

A Service of

ZBW

Leibniz-Informationszentrum Wirtschaft Leibniz Information Centre for Economics

Hyytinen, Ari; Maliranta, Mika

Working Paper When do employees leave their job for entrepreneurship: Evidence from linked employer-employee data

ETLA Discussion Papers, No. 1023

Provided in Cooperation with: The Research Institute of the Finnish Economy (ETLA), Helsinki

Suggested Citation: Hyytinen, Ari; Maliranta, Mika (2006) : When do employees leave their job for entrepreneurship: Evidence from linked employer-employee data, ETLA Discussion Papers, No. 1023, The Research Institute of the Finnish Economy (ETLA), Helsinki

This Version is available at: https://hdl.handle.net/10419/63848

Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

Terms of use:

Documents in EconStor may be saved and copied for your personal and scholarly purposes.

You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.

If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.



WWW.ECONSTOR.EU



Keskusteluaiheita – Discussion papers

No. 1023

Ari Hyytinen* – Mika Maliranta**

WHEN DO EMPLOYEES LEAVE THEIR JOB FOR ENTREPRENEURSHIP: EVIDENCE FROM LINKED EMPLOYER-EMPLOYEE DATA***

* Bank of Finland, PO Box 160, FI-00101 Helsinki E-mail: ari.hyytinen@bof.fi.

** The Research Institute of the Finnish Economy, Lönnrotinkatu 4B, FIN-00120 Helsinki, Finland. E-mail: mika.maliranta@etla.fi.

*** We thank Petri Böckerman as well as conference/seminar participants at The Government Institute for Economic Research (VATT), the ONS Analysis of Enterprise Microdata Conference (CAED 2005) in Cardiff and the XXII Meeting of Finnish Economists in Jyväskylä (2005) for useful comments. This paper has its origins in a research project on entrepreneurship at Etla/Etlatieto, funded by the National Technology Agency of Finland, Tekes (project 579/31/03 and 10582/25/04). Maliranta gratefully ac-knowledges funding within the Research Programme for Advanced Technology Policy (ProAct) by Tekes.

The views expressed are those of the authors and do not necessarily reflect the views of the Bank of Finland or the Research Institute of the Finnish Economy (ETLA).

HYYTINEN, Ari – **MALIRANTA**, Mika, **WHEN DO EMPLOYEES LEAVE THEIR JOB FOR ENTRE-PRENEURSHIP: EVIDENCE FROM LINKED EMPLOYER-EMPLOYEE DATA.** Helsinki: ETLA, Elinkeinoelämän Tutkimuslaitos, The Research Institute of the Finnish Economy, 2006, 24 p. (Keskusteluaiheita, Discussion papers, ISSN 0781-6847; No. 1023).

ABSTRACT: Existing firms are argued to be an important source of new entrepreneurs. Yet, relatively little is known about the characteristics of firms that breed new entrepreneurs. We use a large linked employee-employer dataset to trace and characterize the types of firms from which new entrepreneurs come in Finland. We find evidence for entrepreneurial learning in smaller firms, for they spawn new entrepreneurs more frequently than larger firms. We also find that the productivity of firms and their R&D-intensity are negatively related to the probability that employees transit into entrepreneurship. These results are robust to controlling for a number of employee and employer attributes.

Keywords: entrepreneurship, occupation choice, mobility

JEL-code: G14, G31, G32

HYYTINEN, Ari – **MALIRANTA**, Mika, **MILLOIN PALKANSAAJAT JÄTTÄVÄT TYÖNSÄ YRITTÄJYYDEN VUOKSI: HAVAINTOJA YHDISTETYSTÄ TYÖNANTAJA-TYÖNTEKIJÄ-AINEISTOSTA.** Helsinki: ETLA, Elinkeinoelämän Tutkimuslaitos, The Research Institute of the Finnish Economy, 2006, 24 s. (Keskusteluaiheita, Discussion papers, ISSN 0781-6847; No. 1023).

TIIVISTELMÄ: Toiminnassa olevien yritysten voidaan ajatella olevan uusien yrittäjien tärkeä lähde. Niiden yritysten ominaisuuksista, joiden työntekijöistä tulee uusia yrittäjiä, tiedetään kuitenkin varsin vähän. Käytämme laajaa yhdistettyä työntekijä-työnantaja-aineistoa jäljittääksemme ja kuvataksemme sitä, minkä tyyppisistä yrityksistä uusia yrittäjiä tulee. Löydöksemme kertoo, että pienten yritysten työntekijöistä tulee uusia yrittäjiä todennäköisemmin kuin suurten yritysten työntekijöistä. Tämä voi kertoa mm. siitä, että pienissä yrityksissä työskentelevät oppivat työssään yrittäjyyteen tarvittavia taitoja ja asioita. Havaitsemme myös, että yrityksen tuottavuus ja T&K intensiteetti vähentää henkilökunnan todennäköisyyttä siirtyä yrittäjyyteen. Tuloksemme eivät muutu, vaikka työntekijä- ja työnantaja-ominaisuuksien vaikutuksia otetaan huomioon monin eri tavoin.

Avainsanat: yrittäjyys, uravalinta, liikkuvuus

JEL-luokittelu: G14, G31, G32

Table of Contents

1	INTRODUCTION	1
2	EMPIRICAL FRAMEWORK	4
3	DATA	5
	3.1 DESCRIPTION OF DATA SOURCE AND TRANSITIONS	5
	3.2 DEFINITIONS OF CONDITIONING VARIABLES	7
4	ANALYSIS OF TRANSITIONS	10
	4.1 CONDITIONAL TRANSITIONS	10
	4.2 MNL ESTIMATIONS	10
	4.3 ROBUSTNESS ANALYSIS	13
5	CONCLUSIONS	15
RE	FERENCES	18

1 INTRODUCTION

Existing firms are argued to be an important source of new entrepreneurs, especially in the U.S. (e.g., Bhide 1994, Gompers, Lerner and Scharfstein 2005, Hellman 2005).¹ Yet, little is known about the firms that breed (or "spawn") new entrepreneurs: From what kind of firms do employees departure to become entrepreneurs? In particular, do certain corporate attributes increase labor mobility and particularly the likelihood of an employee leaving her job for entrepreneurship? We take in this paper advantage of a unique and large linked employee-employer dataset from Finland to address these questions. Besides having around 1.5 million person-year observations (0.44 million private sector employees during 5 years), the strength of this dataset is that it allows us i) to trace and characterize the types of firms from which each new entrepreneur comes from and ii) to contrast these transitions into entrepreneurship with other forms of labor market mobility.

The available literature identifies a couple of prominent firm attributes that are likely to have an effect on the rate at which an established company breeds new entrepreneurs. The first of them is firm size. Size matters, because especially smaller incumbents can serve as hatcheries in which entrepreneurial learning takes place. Opportunities for such learning are many, not least because the employees of smaller firms often work alongside of the firms' manager-founder(s), allowing them to observe how small businesses are run (see, e.g., Gompers, Lerner and Scharfstein 2005 and the references therein). A contrasting view of the effect of firm size is that employees are pushed from large, bureaucratic firms into entrepreneurship

¹ Bhide (1994) documents, for example, that 71% of the founders of companies on the 1989 Inc. 500 list of the fastest growing companies in the U.S. essentially replicated or modified an idea that they had encountered in their previous employment.

because of the reluctance of such firms to develop their employees' entrepreneurial ideas further within the firm.²

The second firm attribute that is likely to have an effect on a firm's likelihood to breed new entrepreneurs is its innovativeness. The likelihood that an employee learns about new technologies, product innovations and new forms of organizing production is more likely, the more R&D-intensive and innovative her current employer is. Albeit employees may pay for transferable knowledge (human capital) through lower wages (Møen 2005), the incentive to make commercial use of an innovative employer's knowledge by quitting and starting a rival start-up may remain (Arrow 1962). While this view suggests a direct relation between a firm's innovativeness and the likelihood of the firm spawning new entrepreneurs, there are confounding effects. Kim and Marschke (2005) argue, for example, that the greater the risk that the employees of a firm departure (to join or start a competitor), the eager the firm is to prevent unauthorized use of its knowledge stock through, e.g., patenting. Another protective measure is to increase the firm's ability to capitalize its employees' ideas within the firm (cf. Gromb and Scharfstein 2003; Hellmann 2005).³ These firm-level protective measures may lead to an inverse relation between a firm's innovativeness and the likelihood of the firm spawning new entrepreneurs.

² The organizational capacity of larger - and presumably more bureaucratic and hierarchical - firms to respond to entrepreneurial ideas (and change more generally) may for example be limited (Henderson 1993). Such organizations also process soft information about new business ideas rigidly (Berger et al. 2004; Stein 2002) and have internal capital markets that may disproportionately favor the established lines of business (Scharfstein and Stein 2000). This view is termed by Gompers et al. the "Xerox view" of entrepreneurial spawning, so called because Xerox is (according to Gompers et al.) one of most prominent examples of large bureaucratic firms whose top executives were in the 1960s and 1970s reluctant to fund its employees' entrepreneurial ideas.

³ A firm's ability to capitalize its employees' ideas reflects at least party its willingness to allow for "intrapreneurship", i.e., within-firm entrepreneurship.

Neither theory nor available evidence has pinned down whether the relation between these two prominent corporate attributes - size and innovativeness - and the likelihood that an employee leaves her job for entrepreneurship, is direct or inverse. In this study, we try to uncover the direction of these relations. We do so by investigating occupational choices in the Finnish labor market in 1997-2002 and particularly by tracing the types of firms from which those who transit into entrepreneurship come from.

Using these data, we find evidence for entrepreneurial learning in smaller firms, for they spawn new entrepreneurs more frequently than larger firms. The result is robust to controlling for firm productivity and R&D-intensity, industry, foreign ownership and a number of employee attributes. This finding is not due to the employees of smaller firms being generically more mobile, as in our data, firm size is not similarly (i.e., inversely) related to the likelihood that a private-sector employee switches into a new job. We also document that the relation between a firm's innovativeness and the likelihood of the firm spawning new entrepreneurs is inverse: Both the productivity and R&D-intensity of a firm are negatively related to the probability that one of its employees transits into entrepreneurship. The inverse relation is not due to the employees of less innovative firms being intrinsically more mobile, as our proxies for firm innovativeness have either no or a positive effect on the likelihood of interfirm labor market switches. In sum, our findings allow us both to reject the view that employees are pushed from larger, bureaucratic firms into entrepreneurship and to question the (often-cited) conjecture that especially the most innovative firms are at risk to loose good ideas and employee-inventions because their employees are most prone to quit and rush to establish new firms.

The rest of the paper is organized as follows: In the next section we outline a framework for our empirical analysis. In section 3 we discuss the data. In section 4 we present the results of our empirical analysis. Section 5 contains a brief summary.

2 EMPIRICAL FRAMEWORK

Consider an employee, labeled *e*, who faces a choice among staying in her current job (*Stay*), switching to a new job (*Switch*), becoming an entrepreneur (*Selfemp*) and transiting into unemployment (*Unemp*). The utility she obtains from alternative *j* is U_{ej} , where $j \in \{Stay, Switch, Selfemp, Unemp\}$. The behavioral rule she follows when making her occupational choice is to choose alternative *i* if and only if $U_{ei} \ge U_{ej}$ for all $j \ne i$. Since not all aspects of utility are observed, we set $U_{ej} \equiv V_{ej} + \varepsilon_{ej}$, where V_{ej} is the observed part and ε_{ej} is the unobserved part. We do not observe attributes of the various alternatives, but instead have a vector of employees' characteristics and their current employer's attributes, labeled x_e . The observed part is assumed to be linear in parameters with constant, i.e., $V_{ej} = x_e \, \beta_j + \alpha_j$ and each ε_{ej} is independently and identically distributed and of type I extreme value. With these assumptions, the multinomial logit (MNL) choice probability is (McFadden 1974):

$$\Pr(e's \ choice = i) \equiv P_{ei} = \frac{\exp(x_e'\beta_i + \alpha_i)}{\sum_{j=1}^{4} \exp(x_e'\beta_j + \alpha_j)}$$
(1)

for e = 1, ..., N.

As the attributes of the alternatives are not observed, the parameters of this model are unidentified unless the parameter vector of one of the alternatives is normalized. For estimation we set $\beta_{Stay} = \alpha_{Stay} = 0$, in which case the remaining coefficients measure the change relative to the employees who stay in their current job. Under this normalization, $\alpha_{Selfemp}$ is for example interpreted as the average effect of un-included factors on the utility of becoming self-employed relative to staying in one's current job.

We estimate the model by the method of maximum likelihood and report average marginal effects. The marginal effects are evaluated for each individual and then averaged over the sample, which means that we report $N^{-1}\sum_{e=1}^{N}\beta_i P_{ei}(1-P_{ei})$ (see Cameron and Trivedi 2005 and Bartus 2005). The marginal effects measure the impacts of infinitesimal changes in the continuous variables and discrete changes in the dummy variables. To allow for within-firm correlation in employees' propensity to leave their firm, we use standard errors that are clustered at the level of plants in which the employees work.

A well-known weakness of the MNL model is that the ratio of P_{ei} to P_{ek} does not depend from alternatives other than *i* and *k*. To check that this independence from irrelevant alternatives (IIA) is not what drives our results, we also experiment with a multinomial probit (MNP) model. Albeit the MNP model is flexible enough to allow for any pattern of substitution, it is not a panacea for us: Parameter stability/identification is problematic in the MNP models, particularly in cases like ours, where data on the attributes of alternatives are unavailable (Keane 1992; see also Cameron and Trivedi 2005).

3 DATA

3.1 Description of data source and transitions

The data used in this paper is a random sample from the Finnish Longitudinal Employer-Employee Data (FLEED) of Statistics Finland. The FLEED data are constructed by linking various administrative registers, such as Employment Statistics (in Finnish: Työssäkäyntitilasto), Business Register, Financial Statements Statistics and the R&D survey of Statistics Finland. The basic unit in this dataset is an individual who belongs to the working population of Finland and who – if organizationally employed – can in most cases be linked to the company and plant in which she works. The original FLEED dataset has three characteristics that are particularly important for this study: The dataset (i) follows over time basically the *entire* working population of Finland, (ii) includes a wealth of information about the individuals and their occupations and (iii) makes it possible to trace individuals' labor market transitions.

The sample available to us covers years 1997-2002 and consists of roughly every third individual in the original FLEED data. In this study we restrict our analysis on the labour market behavior of (business) sector employees, which leaves us with about 1.490 million person-year observations that consist of 441 000 individuals during a 5-year period.

Table 1 gives an overview of the data and particularly of the transitions we observe: transitions from salary work into a new job (*Switch*), self-employment (*Selfemp*) and unemployment (*Unempl*). The table also reports the share of immobile employees, who do not leave their current job (*Stay*). As this transition matrix shows, around 20% of those who are employed in year t transit into a new occupation during year t+1. About 15% of employees switch annually to a new job and around 4-5% transit into unemployment. Only 0.6-0.7% of the employed transit into entrepreneurship, which makes these transitions a relatively rare labor market event.

[INSERT TABLE 1 ABOUT HERE]

The transition matrix of Table 1 shows in addition that the shares of individuals who make a move on the labor market have remained constant over time. The stability of these aggregate shares goes well with the finding that also the determinants of individual-level transitions have in our data been stable: As we report in more detail in connection with the robustness tests, the results we are about to report do not depend on the time period we choose. We have, in particular, run all the univariate analyses and MNL estimations of this paper separately for each year in our data, but found no major differences between the years. It is therefore appropriate to focus on a shorter period. In what follows, we concentrate on those individuals who

were employed at the end of 2001 and examine the transitions that they made by the end of 2002. As shown on the last row of Table 1, this restriction means that we investigate occupational choices of 308938 employees. These employees either stayed in their current job (80.09%) or transited from salary work into a new job (15.01%), unemployment (4.14%) or entrepreneurship (0.67%) in 2002. Our aim is to trace the types of firms from which the 2180 employees who transited into entrepreneurship came from and contrast these transitions with the other forms of labor market mobility.

3.2 Definitions of conditioning variables

Firm size and innovativeness

To study how the size of firms for whom the employees in our data worked at the end of year t affect transitions in year t+1, we need a measure for firm size. The measure we use is the logarithm of the number of employees the firm (plant) had at the end of the year t. This variable, labeled *Lnsize*, is based on the midpoint of the seven employment categories to which each firm (plant) in our data has been assigned to. We reverse-engineered the categorical size variable into a continuous one to ease the interpretation of our results, but it is worth stressing already here that our results are robust to using dummy variables instead (see robustness tests for details).

Measuring firm innovativeness is rarely straightforward, which is why we use two proxies. The first, *Lnprod*, is a standard measure of firm productivity and equals the logarithm of the ratio of value added per person. The second, *R&D-dummy* is a dummy that equals one if the ratio of R&D expenditures to turnover exceeds 3.5%, and is zero otherwise.⁴ Because a firm's productivity originates both from informal and formal R&D as well as from process,

⁴ This cutoff level is exogenously given to us, as the data do not allow us to reliably measure the ratio of R&D expenditures to turnover for firms that do only a little formal R&D.

product and organizational innovations, *Lnprod* is a broader measure of a firm's innovativeness than *R&D-dummy*. Our results are robust to using an alternative measure of firm productivity and to not including *R&D-dummy* simultaneously into the models (again, see robustness tests for details).

Individual characteristics

We use a number of control variables to capture the effect of an individual's characteristics on the propensity of her leaving her current employer. These conditioning variables include *Schooling* (= the number of schooling years typically needed for one's highest degree), Age (= age in years), Age2 (= age squared), and *Gender* (= 1 if female and zero otherwise), which a number of previous studies have found to be important determinants of labor market switches. We include *Tenure* (= years of firm-specific experience, as measured by the length of the current employment) and *Tenure2* (= tenure squared). We include these control variables, as the relation between tenure and quit intentions may be negative due to employee heterogeneity (even when there is no true negative state dependence in turnover). Moreover, firms may use wage as a means of lowering the quit rates, which suggests that *Lnincome* (= logarithm of total taxable income in year *t*) should be included. This variable may also proxy the productivity of the employee in her current job, as well as her ability.

As transitions from paid work into entrepreneurship may be intimately linked to the availability of capital, we control for *Wealth* (= 1 if taxable wealth is higher than the 75th percentile in year *t*) and *Spouse's income* (= 1 if the spouse's taxable income in year *t* is higher than the median of such incomes, and = 0 otherwise).⁵ We further control for saving behavior by including *Savings rate* (= taxable wealth at *t* minus taxable wealth at *t*-2 divided by the av-

⁵ Evans and Jovanovic (1989) provide empirical evidence from the U.S. and Johansson (2000) from the Finnish perspective.

erage taxable income during the two-year period). The reason for constructing this variable and using it as a control is that initially wealth constrained people may work hard for someone else and save in the hope of eventually becoming a self-financed entrepreneur (Ghatak, Morelli and Sjöström 2001).

Besides the individual-specific controls listed here, we have experimented with other specifications in which additional individual characteristics have been included in the MNL model. The results of these estimations are reported in connection with robustness tests.

Other firm controls

We use a number of controls for firm attributes other than size and innovativeness. These controls include *Foreign-ownership* (= 1 if foreign-owned at the end of year *t*, and 0 otherwise)⁶, and *Declining employment* (= 1 if a plant's employment shrank between year *t* and *t*-2, and 0 otherwise). We also control for industry (24 categories, based on NACE Rev. 1-classification) and the age of the establishment (7 categories, based on the year of entry) for which the employees work at the end of year *t*.

The mean, standard deviation, as well as 1%, 25%, 50%, 75% and 99% percentiles of each conditioning variable are given in Appendix 1. These statistics pertain to year t, which refers to 2001. The appendix also reports these statistics conditional on the type of labor market transition that takes place in year t+1 (i.e., 2002).

⁶ Ownership status is based on "the ultimate beneficiary owner" (UBO). In our classification, a firm is labeled foreign owned when the ultimate foreign ownership is at least 20 per cent.

4 ANALYSIS OF TRANSITIONS

4.1 Conditional transitions

Table 2 conditions the transitions taking place in year t+1 on firm size and innovativeness. The conditioning variables are *Lnsize* (Panel A, with the seven size categories in the table corresponding to the original employment categories that are available in the data), *Lnprod* (Panel B, with four productivity quartiles) and *R&D-dummy* (Panel C). The panel shows that new entrepreneurs come rarely from the largest or the most productive firms, as the probability of an entry into self-employment is decreasing in firm size and innovativeness. So are switches into unemployment, whereas interfirm switches take most frequently place from the smallest and largest (employing at least 300 persons) firms. The fact that interfirm transitions refer to employee movements between establishments. Therefore, a proportion of the switches into a new job are intra-company transitions in the multi-establishment companies. In our data, 30% of all switches to a new job take place within the same company, whereas the corresponding number for the largest companies is 54%.⁷

[INSERT TABLE 2 ABOUT HERE]

4.2 MNL estimations

Table 3 presents the average marginal effects from the MNL estimations, together with their standard errors and *p*-values. In the first column (Model 1), only industry dummies are included in addition to *Lnsize*, *Lnprod* and *R&D-dummy*. For the second column (Model 2) we include the remaining firm/plant level controls, whereas for the third column (Model 3), we

⁷ A few percent of those employees that have stayed in the same establishment have moved to another company because the owner-company of the establishment has changed during the year.

only add the individual characteristics. The marginal effects in the fourth column (Model 4) are based on a specification that includes both sets of control variables.

The table shows that Pr(choice = Selfemp) is decreasing in *Lnsize*, *Lnprod* and *R&D*dummy in each column and that the magnitudes of the estimated effects are stable across the columns. These results reinforce the earlier univariate finding of ours that both firm size and innovativeness decrease the likelihood of a firm spawning new entrepreneurs. Moreover, it seems that the negative effect of firm size and innovativeness on entrepreneurial spawning is not due to the employees of smaller and more innovative firms being generically more mobile: In columns 1-4, *Lnsize* is directly related to Pr(choice = Switch), whereas the effect of *Lnprod* on the probability is unstable (with a positive effect in columns 3 and 4). The effects of *Lnsize* and *Lnprod* on Pr(choice = Switch) are similar to these variables effects on Pr(choice = Selfemp) even if we allow for non-linearities (see robustness tests). Finally, *R&D-dummy* has a positive and statistically significant effect on Pr(choice = Switch). This finding is in contrast to how *R&D-dummy* affects Pr(choice = Selfemp).

[INSERT TABLE 3 ABOUT HERE]

The results of our MNL estimations speak for entrepreneurial learning in smaller firms, for they spawn new entrepreneurs more frequently than larger firms. This result is consistent with the view that the employees of smaller firms learn how small businesses are run when working alongside of the firms' manager-founder(s). The negative relation can also emerge because of a number of related reasons: First, the employees of smaller firms may learn a balanced set of skills necessary to start up a business (Lazear 2004), because there are fewer opportunities for within-firm specialization in such firms. Second, it is likely that in a small

firm, the job tasks that employees are assigned are frequently multi-faceted. The employees of smaller firms are therefore often exposed simultaneously to a network of customers and suppliers of labor, production technologies and capital (Saxenian 1994). Such exposure certainly eases an employee's transition into entrepreneurship.

It is worth emphasizing that the negative relation between firm size and entrepreneurial spawning is robust to simultaneously controlling for firm productivity and (past) growth of the spawning firms. This suggests that it is not the innovativeness of smaller firms either that encourages spawning.

Our results are not consistent with the view that innovative/R&D-intensive firms are particularly prone to breed new firms. Instead, the results are consistent with such firms taking protective measures against extensive employee transition into entrepreneurship. Examples of such measures are use of patenting (Kim and Marschke 2005) to discourage employee start-ups and active encouragement of within-firm exploitation of employees' ideas (Gromb and Scharfstein 2003). The positive marginal effect of R&D-dummy on Pr(choice = Switch) is, however, consistent with the existence of *interfirm* transfers of business and technological knowledge embodied in the human capital of the switching employees (Arrow 1962, Møen 2005, Kim and Marschke 2005).

A final observation that emerges from Table 3 is that the probability of an entry into unemployment is decreasing in *Lnsize*, *Lnprod* and *R&D-dummy*.⁸ These effects are similar to what we found for the likelihood that an employee becomes self-employed. Does this mean that self-employment is unemployment in disguise (cf. Earle and Sakova 2000)? It might

⁸ To save space, we do not report in the table the marginal effects of the control variables, but note just briefly that the data provide some evidence for a wealth -effect: The propensity of becoming an entrepreneur is increasing in *Wealth* and *Spouse's income*. The average marginal effect of *Savings rate* is positive, but insignificant (*p*-value = 0.12).

mean, if the similarity of the effects emerges because smaller and less innovative firms *push* employees out. However, it is not clear why that effect would not be captured by *Declining employment* -variable (and other controls) nor why the push would show up as a higher likelihood of transiting into un- and self-employment *but not* as a higher likelihood of interfirm mobility.

4.3 Robustness analysis

We have run a number of robustness tests:

First, we have repeated all the univariate analyses and MNL estimations separately for each year in our data. Our results are not dependent on the time period we choose: For example, when we reproduce Table 2 separately for t = 1997, ..., 2000, the differences to Table 2 (with t = 2001) are minor. What's more, yearly MNL estimations using data from these earlier periods produce results that are very similar to those reported in Table 3 (these results are available from the authors on request).

Second, to check that deficiently measured firm productivity, ignoring the efficiency in the usage of capital input, is not what drives our results we change *Lnprod* to a measure of firm-level total factor productivity. None of the results in Table 3 are challenged when this alternative measure of productivity is used: The average marginal effect of the logarithm of the total factor productivity on Pr(choice = Selfemp) is for example -0.00689 (*p*-value < 1%). Moreover, if we drop *R&D-dummy*, the results for *Lnprod* become stronger. Finally, if we drop *Lnprod*, firm size still has a negative effect on Pr(choice = Selfemp).

Third, Table 2 suggested that the effect of *Lnsize* and *Lnprod* on Pr(choice = Switch) might be non-linear. We can allow for such effects in the MNL estimations by including squared terms for both variables: None of our basic findings change as a result. In particular, we again find a negative (but decreasing) effect of *Lnsize* and *Lnprod* on

Pr(choice = Selfemp) over the relevant range. Moreover, switches into unemployment or to a new job are not similarly related to the size of firms as are the transitions into entrepreneurship. If anything, switching to a new job is more likely for an employee working in larger firms. Neither has firm productivity the same, uniformly negative effect on Pr(choice = Switch) as it has on Pr(choice = Selfemp) when these more flexible specifications are used.

Fourth, the results of columns 3-4 (Models 3 and 4) show that the negative effects of *Lnsize*, *Lnprod* and *R&D-dummy* on Pr(choice = Selfemp) are robust to controlling for a number of employee attributes, such as schooling, age, gender, tenure, wage (income), taxable wealth, spouse's income, and ability to save. As many of these controls are often-used proxies for a person's ability, we have no reason to expect that employees' ability, possible correlated with certain firm characteristics, would be a major driver of our findings. However, we can do a bit more by demonstrating that the MNL estimations are robust to adding a number of new control variables into the model. The variables we have tried include years of schooling of the co-workers (proxied by the highest degree achieved), employee's tenure choice (i.e. homeownership), and a variable capturing falling average income in the plant in which the employee works. We also experimented by splitting the sample into three groups, the first for those with low education (secondary level education or lower), the second for those with medium level education (lower-degree level tertiary education) and the third for those with the university level education. This robustness test addresses the concern that most of employee transitions into self-employment have little to do with innovative activity (and thereby with the innovativeness of present employer), especially in the case of less educated workers. These estimations showed that the more innovative firms do not breed new entrepreneurs

more frequently than the less innovative firms even if we focus on the group of better educated private sector employees.⁹

Fifth, to check whether the IIA property of the MNL model is driving our results we have done two things: i) we dropped *Unemp* from the choice set and re-estimated the models reported in Table 3. The main results remain unchanged; ii) we have estimated a MNP model. Despite its deficiencies in cases where there are no data on the attributes of alternatives (Keane 1992), the MNP estimations confirms the basic findings of Table 3.

Finally, the MNL results that we have presented so far are based on standard errors that are clustered at the level of plants. Our results are robust to not clustering at all, or clustering at the firm-level.

5 CONCLUSIONS

This paper documents that during recent years, 0.6-0.7% of the private sector employees have annually transited into entrepreneurship in Finland. As compared to other types of labour market transitions, the transitions into entrepreneurship are a relatively rare labor market event.

As far as we know, the only pieces of available evidence on the characteristics of firms that breed new entrepreneurs come from Gompers, Lerner and Scharfstein (2005). Using a sample of 1370 public companies from the U.S., they find that especially young firms that were once innovative and venture capital -backed and are located in the U.S. hotbeds of

⁹ More specifically, the results (available on request) for those with the university level education indicate that both the size and the innovativeness of the employer (measured by productivity and R&D intensity) have a significant positive effect on the propensity to switch the employer whereas these factors significantly decrease an individual's probability to leave her job for entrepreneurship. The effect of size and innovativeness on the propensity to switch to a new job is hence different from their effect on the propensity to transit intro entrepreneurship. The difference seems to be more pronounced for the highly educated workers.

venture capital activity (such as Silicon Valley) are important sources of new firms. Inspired in part by this finding, the aim of this paper has been to shed new light on these characteristics by focusing on the role of firm size and innovativeness.

Our findings corroborate the main result of Gompers et al., but only partially:

- Like Gompers et al., we find evidence for entrepreneurial learning in smaller, entrepreneurial firms, for they spawn in our data new entrepreneurs more frequently than larger firms;
- Unlike like Gompers et al., we document an inverse relation between the innovativeness of firms and the likelihood of them spawning new entrepreneurs.

In sum, the results of this paper allow us to reject the view that employees are pushed from larger, bureaucratic firms into entrepreneurship. Nor can we bring evidence for the often cited conjecture that especially the most innovative firms are most likely to loose good ideas when their employees quit and rush to establish new firms. Whether this finding emerges because such firms are wittingly (and capable) protecting their IPRs or because they are good at internally capitalizing their employees' ideas is an interesting question that clearly warrants further research.

Our study does not answer the question of whether, and if so by which means, we should encourage employees to leave their job for entrepreneurship, especially if they work for an innovative, private sector employer. Yet, because the pool of current business sector employees is often perceived to constitute the most prominent source of high-quality entrepreneurship in developed economies, the finding that transitions into entrepreneurship are relatively rare and that the innovative incumbent companies are *not* the primary sources of

new entrepreneurship sound somewhat disturbing, if not alarming.¹⁰ The reason for our concern is that a growing theoretical and empirical literature emphasizes the process of "creative destruction" as a source of long-term economic growth (see e.g. Foster, Haltiwanger, Krizan 2001; Klette and Kortum 2004). The process needs to be incessantly nourished by innovations and market experimentation of new ideas. Experimentation, in turn, calls for a sufficient supply of high-quality, daring entrepreneurs. The most recent analyses suggest that these related market processes, i.e., selection of talent into entrepreneurship and market experimentation, are especially instrumental for the long-term growth of the economies close to the global technology frontier (see, e.g., Acemoglu, Aghion and Zilibotti 2006).

¹⁰ On the other hand, new innovation-induced entry is only one side of the long-term growth process. Successful adoption and commercialization of innovations within existing firms is yet another source of growth. It is likely to call for a considerable amount of high-quality "intrapreneurship" among employees; see also Holmes and Schmitz (1990).

References

- Acomoglu, Daron, Aghion, Philippo and Zilibotti, Fabrizio 2006, Distance to Frontier, Selection, and Economic Growth, *Journal of the European Economic Association*, Vol, 4, Issue 1, pp. 37-74.
- Arrow, Kenneth, 1962, Economic welfare and the allocation of resources for invention, in: *The Rate and Direction of Inventive Activity: Economic and Social Factors* (editor R. R.Nelson), NBER Special Conference Series, Vol. 13, Princeton.
- Bartus, Tamás, 2005, Estimation of marginal effects using margeff, Stata Journal, Vol. 5, pp. 309-329.
- Berger, Allen N., Miller, Nathan H., Petersen, Micthel A., Raghuram G. Rajan, and Stein, Jeremy C., 2004, Does function follow organizational form? Evidence from the lending practices of large and small banks, *Journal of Financial Economics*, to appear.
- Bhide, Amar V., 1994, How entrepreneurs craft strategies that work, *Harvard Business Review*, March-April 1994, pp. 50-161.
- Cameron, A. Colin and Trivedi, Pravin K., 2005, *Microeconometrics: Methods and Applications*, Cambridge: Cambridge University Press.
- Earle, John S. and Sakova, Zuzana, 2000, Business start-ups or disguised unemployment? Evidence on the character of self-employment from transition countries, *Labour Economics*, Vol. 7, pp. 575-601.
- Evans, David S. and Jovanovic, Byan, 1989, An estimated model of entrepreneurial choice under liquidity constraints, *Journal of Political Economy*, Vol. 97, No. 4, pp. 808-827.
- Foster, Lucia, Haltiwanger, John and Krizan C.J., Aggregate productivity growth: Lessons from microeconomic evidence, in *New developments in productivity analysis* (editors Hulten, Charles R., Dean, Edwin R. and Harper, Michael J.), NBER Studies in Income and Wealth, Vol. 63, University of Chicago Press
- Ghatak, Maitreesh, Morelli, Massimo, and Sjöström, Tomas, 2001, Occupational choice and dynamic incentives, *Review of Economic Studies*, Vol. 68., pp. 781-810.
- Gompers, Paul, Lerner, Josh, and Scharfstein, David, 2005, Entrepreneurial spawning: Public corpotrations and the genesis of new ventures, 1986-199, *Journal of Finance*, Vol. 60, No. 2, pp. 577-615.
- Gromb, Denis and Scharfstein, David, 2003, Entrepreneurship in equilibrium, mimeo, Harvard Business School.
- Hellmann, Thomas, 2005, When do employees become entrepreneurs? Mimeo, University of British Columbia.
- Henderson, Rebecca, 1993, Underinvestment and incometence as responses to radical innovation: Evidence from the photolithographic alignment equipment industry, *Rand Journal of Economics*, Vol. 24, pp. 248-270.

Holmes, T. J. and Schmitz Jr., J. A., 1990, A theory of entrepreneurship and its application to the study of business transfers. Journal of Political Economy, Vol. 98, Issue 2, 265-294.

- Johansson, Edvard, 2000, Self-employment and liquidity constraints: Evidence from Finland, *Scandinavian Journal of Economics*, Vol. 102, pp. 123-134.
- Keane, Michael P., 1992, A note on identification in the multinomial probit model, *Journal of Business and Econoic Statistics*, Vol. 10. No. 2, pp. 193-200.

- Kim, Jinyoung and Marschke, Gerald, 2005, Labor mobility of scientists, technological diffusion, and the firm's patenting decision, *Rand Journal of Economics*, Vol. 36, Issue 2, pp. 298-317.
- Klette, Tor and Kortum, Samuel, 2004, Innovating firms and aggregate innovation, *Journal of Political Economy*, Vol. 112, Issue 5, pp. 986-1018.
- Lazear, Edvard P., 2004, Balanced skills and entrepreneurship, *American Economic Review* (Papers and Proceedings), Vol. 94, No. 2, pp. 208-211.
- McFadden, Daniel, 1974, Conditional logit analysis of qualitative choice behaviour, in: P. Zarembka (ed.), *Frontiers in Econometrics*, Academic Press, New York, pp. 105-142.
- Møen, Jarle, 2005, Is mobility of technical personnel a source of R&D spillovers? *Journal of Labor Economics*, Vol. 23, No. 1, pp. 81-114.
- Saxenian, AnnaLee, 1994, Regional Advantage: Culture and Competition in Silicon Valley and Route 128, Harvard University Press, Cambridge, MA.
- Scharfstein, David and Stein; Jeremy C., 2000, The dark side of internal capital markets: Divisional rent-seeking and inefficient investment, *Journal of Finance*, December 2000, pp. 2537-2564.
- Stein, Jeremy C., 2002, Information production and capital allocation: Decentralized vs. hierarchical firms, *Journal of Finance*, Vol. 57, October 2002, pp. 1891-1921.

Appendix 1. Descriptive statistics

	Variable	NI	Maar	614	n4	n05	Percentile	n75	n00
	Variable	N	Mean	Std	р1	p25	p50	p75	p99
Stay	Lnsize	247428	4.40	1.76	0.92	2.67	5.30	5.99	5.99
	Lnprod	247428	10.77	0.57	9.37	10.43	10.74	11.05	12.16
	R&D-dummy	247428	0.09	0.28	0.00	0.00	0.00	0.00	1.00
	Schooling years	247428	12.15	2.28	9.00	12.00	12.00	14.00	17.00
	Age	247428	39.97	10.92	19.00	31.00	40.00	49.00	61.00
	Tenure	247428	8.82	9.20	0.00	1.67	5.33	13.00	35.75
	Lnincome	247428	10.11	0.51	8.47	9.88	10.13	10.40	11.31
	Gender	247428	0.36	0.48	0.00	0.00	0.00	1.00	1.00
	Declining employment	247428	0.38	0.49	0.00	0.00	0.00	1.00	1.00
	Foreign-owned	247428	0.17	0.38	0.00	0.00	0.00	0.00	1.00
	Savings rate	247428	0.04	0.12	-0.12	0.00	0.00	0.04	0.31
	Wealth	247428	0.27	0.44	0.00	0.00	0.00	1.00	1.00
	Spouse's income	247428	0.53	0.50	0.00	0.00	1.00	1.00	1.00
Switch	Lnsize	45348	4.30	1.85	0.92	2.67	5.30	5.99	5.99
	Lnprod	45348	10.69	0.74	8.69	10.24	10.62	11.03	12.16
	, R&D-dummy	45348	0.13	0.34	0.00	0.00	0.00	0.00	1.00
	Schooling years	45348	12.39	2.34	9.00	12.00	12.00	14.00	17.00
	Age	45348	34.61	10.97	18.00	25.00	33.00	43.00	59.00
	Tenure	45348	4.74	6.92	0.00	0.58	1.92	5.67	31.67
	Lnincome	45348	9.88	0.72	7.54	9.55	9.99	10.33	11.31
	Gender	45348	0.41	0.49	0.00	0.00	0.00	1.00	1.00
	Declining employment	45348	0.35	0.48	0.00	0.00	0.00	1.00	1.00
	Foreign-owned	45348	0.17	0.37	0.00	0.00	0.00	0.00	1.00
	Savings rate	45348	0.03	0.10	-0.12	0.00	0.00	0.00	0.31
	Wealth	45348	0.17	0.37	0.00	0.00	0.00	0.00	1.00
	Spouse's income	45348	0.46	0.50	0.00	0.00	0.00	1.00	1.00
Calfama			2.74						
Selfemp	Lnsize	2180		1.85	0.92	0.92	1.95	4.31	5.99
	Lnprod	2180	10.50	0.59	8.28	10.24	10.51	10.81	12.05
	R&D-dummy	2180	0.03	0.17	0.00	0.00	0.00	0.00	1.00
	Co-worker schooling	2180	12.07	2.23	9.00	12.00	12.00	14.00	17.00
	Age	2180	37.96	9.67	20.00	31.00	37.00	45.00	59.00
	Tenure	2180	3.96	5.07	0.00	0.75	1.92	5.25	24.25
	Lnincome	2180	9.85	0.73	7.39	9.51	9.92	10.28	11.31
	Gender	2180	0.28	0.45	0.00	0.00	0.00	1.00	1.00
	Declining employment	2180	0.35	0.48	0.00	0.00	0.00	1.00	1.00
	Foreign-owned	2180	0.07	0.25	0.00	0.00	0.00	0.00	1.00
	Savings rate	2180	0.04	0.13	-0.12	0.00	0.00	0.02	0.31
	Wealth	2180	0.24	0.43	0.00	0.00	0.00	0.00	1.00
	Spouse's income	2180	0.54	0.50	0.00	0.00	1.00	1.00	1.00
Jnemp	Lnsize	13982	3.83	1.88	0.92	1.95	3.54	5.99	5.99
	Lnprod	13982	10.54	0.57	8.69	10.25	10.55	10.83	12.16
	R&D-dummy	13982	0.05	0.23	0.00	0.00	0.00	0.00	1.00
	Schooling years	13982	11.39	2.01	9.00	9.00	12.00	12.00	17.00
	Age	13982	40.05	12.65	18.00	29.00	40.00	52.00	62.00
	Tenure	13982	5.20	8.49	0.00	0.50	1.33	5.75	37.08
	Lnincome	13982	9.78	0.61	7.63	9.53	9.88	10.15	10.92
	Gender	13982	0.35	0.48	0.00	0.00	0.00	1.00	1.00
	Declining employment	13982	0.37	0.48	0.00	0.00	0.00	1.00	1.00
	Foreign-owned	13982	0.14	0.35	0.00	0.00	0.00	0.00	1.00
	Savings rate	13982	0.03	0.11	-0.12	0.00	0.00	0.01	0.31
	Wealth	13982	0.22	0.41	0.00	0.00	0.00	0.00	1.00
	Spouse's income	13982	0.39	0.49	0.00	0.00	0.00	1.00	1.00
All	Variable	N	Mean	Std	p1	p25	p50	p75	p99
	Lnsize	308938	4.35	1.79	0.92	2.67	5.30	5.99	5.99
	Lnprod	308938	10.74	0.60	9.31	10.41	10.72	11.04	12.16
	R&D-dummy	308938	0.09	0.29	0.00	0.00	0.00	0.00	1.00
	Schooling years	308938	12.15	2.29	9.00	12.00	12.00	14.00	17.00
	Age	308938	39.17	11.16	19.00	30.00	39.00	48.00	61.00
	Tenure	308938	8.02	8.99	0.00	1.25	4.42	12.08	35.50
	Lnincome	308938	10.06	0.56	8.19	9.83	10.10	10.38	11.30
	Gender	308938	0.36	0.48	0.00	0.00	0.00	1.00	1.00
	Declining employment	308938	0.38	0.48	0.00	0.00	0.00	1.00	1.00
		200020	0.17	0.38	0.00	0.00	0.00	0.00	1.00
	Foreign-owned	308938	0.17	0.00	0.00	0.00		0.00	1.00
	Foreign-owned Savings rate	308938	0.04	0.11	-0.12	0.00	0.00	0.03	0.31
	5								

Year t	Stay	Switch	Selfemp	Unemp	Total
1997	231,964	40,378	1,750	11,169	285,261
	81.32	14.15	0.61	3.92	100.00%
1998	239,949	45,479	1,997	12,925	300,350
	79.89	15.14	0.66	4.30	100.00%
1999	234,968	45,805	2,021	10,371	293,165
	80.15	15.62	0.69	3.54	100.00%
2000	240,471	46,680	2,048	13,299	302,498
	79.50	15.43	0.68	4.40	100.00%
2001	247,428	45,348	2,180	13,982	308,938
	80.09	14.68	0.71	4.53	100.00%
Total	1,194,780	223,690	9,996	61,746	1,490,212
	80.18	15.01	0.67	4.14	100.00%

Table 1: Transition matrix 1997-2001

Table 2: Conditional transitions

Panel A Size of the firm	Stay	Status Switch	in 2001 Selfemp	Unemp	Total
0-4 employees	23,757 73.71	5,425 16.83	793 2.46	2,254 6.99	32,229 100.00%
5-9 employees	18,534	3,766	332	1,443	24,075
	76.98	15.64	1.38	5.99	100.00%
10-19 employees	21,287	3,818	257	1,573	26,935
	79.03	14.17	0.95	5.84	100.00%
20-49 employees	28,830	5,117	235	1,906	36,088
	79.89	14.18	0.65	5.28	100.00%
50-99 employees	20,191	3,202	113	1,182	24,688
	81.78	12.97	0.46	4.79	100.00%
100-299 employees	32,855	4,793	150	1,549	39,347
	83.50	12.18	0.38	3.94	100.00%
300- employees	101,974	19,227	300	4,075	125,576
	81.21	15.31	0.24	3.25	100.00%
Total	247,428	45,348	2,180	13,982	308,938
	80.09	14.68	0.71	4.53	100.00%

Table 2: (continued)

Panel B					
Productivity	Stay	Status Switch	in 2001 Selfemp	Unemp	Total
Quartile 1	46,532 70.91	13,890 21.17	749 1.14	4,454 6.79	65,625 100.00%
Quartile 2	64,274 80.33	10,718 13.40	688 0.86	4,332 5.41	80,012 100.00%
Quartile 3	69,158 84.46	9,116 11.13	443 0.54	3,166 3.87	81,883 100.00%
Quartile 4	67,464 82.86	11,624 14.28	300 0.37	2,030 2.49	81,418 100.00%
Total	247,428 80.09	45,348 14.68	2,180 0.71	13,982 4.53	308,938 100.00%
Panel C	I	Ctotuc	in 2001		
R&D-intensity	Stay	Switch	Selfemp	Unemp	Total
R&D-dummy = 0	226,236 80.54	39,325 14.00	2,118 0.75	13,229 4.71	280,908 100.00%
R&D-dummy = 1	21,192 75.60	6,023 21.49	62 0.22	753 2.69	28,030 100.00%
Total	247,428 80.09	45,348 14.68	2,180 0.71	13,982 4.53	308,938 100.00%

Table 3: MNL estimations

	Model 1		Model 2			Model 3			Model 4			
	Coef.	Std.err.	p-value	Coef.	Std.err.	p-value	Coef.	Std.err.	p-value	Coef.	Std.err.	p-value
Switch												
Lnsize	1.00	0.12	0.00	1.24	0.14	0.00	1.57	0.12	0.00	1.72	0.14	0.00
Lnprod	-1.04	0.54	0.06	-0.60	0.52	0.25	0.38	0.50	0.45	0.67	0.49	0.17
R&D-dummy	6.38	1.37	0.00	6.56	1.36	0.00	6.03	1.39	0.00	6.25	1.38	0.00
Selfemp												
Lnsize	-0.28	0.01	0.00	-0.27	0.01	0.00	-0.26	0.01	0.00	-0.26	0.01	0.00
Lnprod	-0.20	0.03	0.00	-0.20	0.03	0.00	-0.21	0.03	0.00	-0.21	0.03	0.00
R&D-dummy	-0.31	0.11	0.01	-0.31	0.11	0.01	-0.31	0.11	0.01	-0.32	0.11	0.00
Unemp												
Lnsize	-0.22	0.04	0.00	-0.26	0.04	0.00	-0.04	0.04	0.37	-0.08	0.05	0.07
Lnprod	-1.53	0.19	0.00	-1.56	0.19	0.00	-1.08	0.19	0.00	-1.12	0.19	0.00
R&D-dummy	-2.62	0.55	0.00	-2.60	0.55	0.00	-1.78	0.54	0.00	-1.75	0.54	0.00
Industry dummies		Yes			Yes			Yes			Yes	
Firm/plant controls	No		Yes		No		Yes					
Individual characteristics No		No		Yes		Yes						
Number of observations 308938			308938			308938			308938			
Log likelihood	-188026.0		-186626.8		-178838.8			-178054.4				
Pseudo R2		0.041			0.048			0.088			0.092	

Note: Coefficients refer to the average marginal effects. They and the standard errors have been multiplied by 100.

ELINKEINOELÄMÄN TUTKIMUSLAITOS (ETLA)

THE RESEARCH INSTITUTE OF THE FINNISH ECONOMY LÖNNROTINKATU 4 B, FIN-00120 HELSINKI

> Puh./Tel. (09) 609 900 Int. 358-9-609 900 http://www.etla.fi

Telefax (09) 601753 Int. 358-9-601 753

KESKUSTELUAIHEITA - DISCUSSION PAPERS ISSN 0781-6847

Julkaisut ovat saatavissa elektronisessa muodossa internet-osoitteessa: http://www.etla.fi/finnish/research/publications/searchengine

- No 989 MIKA WIDGRÉN, Revealed Comparative Advantage in the Internal Market. 01.08.2005. 20 p.
- No 990 ARI HYYTINEN MIKA PAJARINEN, Yrittäjäksi ryhtyminen ja yrittäjyysasenteet Suomessa: Havaintoja kyselytutkimuksista. 10.08.2005. 19 s.
- No 991 CHRISTOPHER PALMBERG MIKA PAJARINEN, Alliance Capitalism and the Internationalisation of Finnish Firms. 01.11.2005. 39 p.
- No 992 ELIAS OIKARINEN, Is Housing Overvalued in the Helsinki Metropolitan Area? 29.09.2005. 33 p.
- No 993 MIKA MALIRANTA PEKKA ILMAKUNNAS, Decomposing Productivity and Wage Effects of Intra-Establishment Labor Restructuring. 02.11.2005. 26 p.
- No 994 VILLE KAITILA MAARIT LINDSTRÖM EWA BALCEROWICZ, Puolan liiketoimintaympäristö ja suomalaisten yritysten kokemukset. 10.11.2005. 72 s.
- No 995 SERGEY SUTYRIN VLADIMIR SHEROV, Russian Regions and Their Foreign Trade. 25.11.2005. 26 p.
- No 996 HANNU PIEKKOLA, Public Funding of R&D and Growth: Firm-Level Evidence from Finland. 20.12.2005. 30 p.
- No 997 AIJA LEIPONEN, Clubs and Standards: The Role of Industry Consortia in Standardization of Wireless Telecommunications. 08.12.2005. 44 p.
- No 998 EWA BALCEROWICZ, Poland's Enterprise Environment A Polish View. 10.01.2006. 19 p.
- No 999 STEFAN NAPEL MIKA WIDGRÉN, The European Commission Appointment, Preferences, and Institutional Relations. 17.01.2006. 20 p.
- No 1000 JUKKA LASSILA TARMO VALKONEN, The Finnish Pension Reform of 2005. 20.01.2006. 20 p.
- No 1001 OLLI-PEKKA OKSANEN, Are Foreign Investments Replacing Domestic Investments? Evidence from Finnish Manufacturing. 19.01.2006. 59 p.
- No 1002 ARTO SEPPÄ, Open Source in Finnish Software Companies. 25.01.2006. 36 p.
- No 1003 TERTTU LUUKKONEN, Venture Capital Industry in Finland Country Report for the Venture Fun Project. 27.02.2006. 48 p.
- No 1004 ELIAS OIKARINEN, Price Linkages Between Stock, Bond and Housing Markets Evidence from Finnish Data. 15.02.2006. 36 p.

- No 1005 JUHA ALHO NIKU MÄÄTTÄNEN, Aggregate Mortality Risk and The Insurance Value of Annuities. 21.02.2006. 15 p.
- No 1006 MORRIS TEUBAL TERTTU LUUKKONEN, Venture Capital Industries and Policies: Some Cross-country Comparisons. 28.02.2006. 23 p.
- No 1007 MIKA PAJARINEN PEKKA YLÄ-ANTTILA, Omistajuus ja yritysten menestyminen: Analyysia suomalaisella aineistolla. 01.03.2006. 42 s.
- No 1008 KARI E.O. ALHO, Labour Market Institutions and the Effectiveness of Tax and Benefit Policies in Enhancing Employment: A General Equilibrium Analysis. 29.03.2006. 43 p.
- No 1010 FRANCESCO DAVERI MIKA MALIRANTA, Age, Technology and Labour Costs. 24.03.2006. 48 p.
- No 1011 MARKKU KOTILAINEN, Economic Shocks, Progressiveness of Taxation, and Indexation of Taxes and Public Expenditure in EMU. 03.04.2006. 29 p.
- No 1012 HELI KOSKI TOBIAS KRETSCHMER, Innovation and Dominant Design in Mobile Telephony. 03.04.2006. 31 p.
- No 1013 HANNU HERNESNIEMI MARTTI KULVIK, Helsingin seudun klusterit sekä erikoistuminen bioteknologiaan ja logistiikkaan. 11.04.2006. 44 s.
- No 1014 LAURA VALKONEN, Deregulation as a Means to Increase Competition and Productivity. Some Finnish experiences. 25.04.2006. 84 p.
- No 1015 VILLE KAITILA, Productivity, Hours Worked, and Tax/Benefit Systems in Europe and Beyond. 27.04.2006. 34 p.
- No 1016 OLAVI RANTALA, Sosiaalietuuksien rahoituksen hinta- ja hyvinvointivaikutukset kotitaloussektorissa. 05.05.2006. 21 s.
- No 1017 MAARIT LINDSTRÖM MIKA PAJARINEN, The Use of Design in Finnish Manufacturing Firms. 05.05.2006. 26 p.
- No 1018 NIKU MÄÄTTÄNEN, Vapaaehtoiset eläkevakuutukset, verotus ja eläkkeelle siirtyminen. 05.05.2006. 25 s.
- No 1019 ESA VIITAMO HANNU HERNESNIEMI, Ympäristöliiketoiminnan määrittely ja tilastollinen seuranta – Ympäristöalalle lisää kilpailukykyä. 15.05.2006. 58 s.
- No 1020 CHRISTOPHER PALMBERG TUOMO NIKULAINEN, Industrial Renewal and Growth Through Nanotechnology? An overview with focus on Finland. 17.05.2006. 45 p.
- No 1021 ESA VIITAMO, Markkinoiden toimivuuden arvioiminen Suuntaviivoja vertailevalle kilpailututkimukselle.
- No 1022 OLLI MARTIKAINEN JUSSI AUTERE MARKKU NURMELA, Performance Improvement in Public Organizations – How to leverage ICT Investments. 30.05.2006. 38 p.
- No 1023 ARI HYYTINEN MIKA MALIRANTA, When do Employees Leave Their Job for Entrepreneurship: Evidence from Linked Employer-employee Data. 31.05.2006. 24 p.

Elinkeinoelämän Tutkimuslaitoksen julkaisemat "Keskusteluaiheet" ovat raportteja alustavista tutkimustuloksista ja väliraportteja tekeillä olevista tutkimuksista. Tässä sarjassa julkaistuja monisteita on mahdollista ostaa Taloustieto Oy:stä kopiointi- ja toimituskuluja vastaavaan hintaan.

Papers in this series are reports on preliminary research results and on studies in progress. They are sold by Taloustieto Oy for a nominal fee covering copying and postage costs.