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Working Paper The global production frontier of universities

KOF Working Papers, No. 272

Provided in Cooperation with: KOF Swiss Economic Institute, ETH Zurich

Suggested Citation: Bolli, Thomas (2011) : The global production frontier of universities, KOF Working Papers, No. 272, ETH Zurich, KOF Swiss Economic Institute, Zurich, https://doi.org/10.3929/ethz-a-006341664

This Version is available at: https://hdl.handle.net/10419/50419

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The Global Production Frontier of Universities

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The Global Production Frontier of Universities**

Thomas Bolli*

February 2011

Abstract

This paper provides first micro-level evidence of the global university production frontier, allowing to estimate technical efficiencies of 273 top research universities across 29 countries between 2007 and 2009. Exploiting comparable international data improves the estimation of the production technology, allows to assess the distance of individual countries to the global frontier and enables comparison of university efficiencies between and across countries. The estimated input distance function uses undergraduate students, graduate students and citations to capture university outputs and staff to measure inputs. Contrasting two alternative econometric strategies to identify technical efficiency yields relatively stable results. Furthermore, the paper addresses the problem of unobserved heterogeneity by relating the obtained efficiency rankings to quality rankings and by exploiting the panel structure of the data to account for unobserved heterogeneity explicitly. The results suggest that technical efficiency rankings can be obtained in a relatively simple econometric setting.

Key words: University, Global Frontier, Efficiency, Stochastic Frontier, Unobserved Heterogeneity, True Random Effects Stochastic Frontier.

JEL Classification: D20, I20

**The author would like to thank the QS World University Rankings for providing the data, Marius Ley, Matthias Bannert and Benjamin Wohlwend for help in preparing it and the participants of the KOF Brown Bag Seminar as well as Spyros Arvanitis, Mehdi Farsi, Marius Ley, Tobias Stucki and Martin Woerter for comments and discussions.

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1 Introduction

The globalization of teaching, research and innovation activities and the corresponding internationalization of the academic world has sparked an increasing demand for comparisons of university quality across countries. The appearance of the QS World University Rankings (QS, 2010) and the Academic Ranking of World Universities (ARWU, 2010) represent the most famous responses to these developments. These two rankings have received substantial interest of the public. They also created a new literature strand that criticizes the employed methodology (see, e.g., Dill and Soo, 2005; Marginson and van der Wende, 2007; Stolz et al., 2010). Furthermore, the rankings evoked a discussion concerning the incentives they create (see, e.g., Hazelkorn, 2007; Clarke, 2007).

However, quality of universities measured by these rankings reflects only one dimension relevant to politicians. The other side of the coin shows the productivity and efficiency in the production process of universities. Hence, complementing these rankings, the literature on university efficiency has grown rapidly as Agasisti and Johnes (2010) and Johnes and Johnes (2009) demonstrate. Worthington (2001) and Johnes (2004) provide literature reviews. However, only few studies provide cross-country evidence and no global production frontier has been established yet (Agasisti and Johnes, 2009).

This paper applies the production function framework to an international data set provided by the QS World University Rankings (QS, 2010) to estimate a multi-output input distance function for 273 top research universities in 29 countries between 2007 and 2009. Thereby the paper extends the existing literature by providing first microlevel evidence of the global university production frontier.

Estimating a global frontier as opposed to individual national production frontiers allows to estimate a more general production technology and to assess the distance of country frontiers to the global frontier. Finally, it allows to compare technical efficiencies between and across countries. However, international comparisons faces problems of data consistency and sample homogeneity (Salerno, 2003). The employed data set circumvents these pitfalls due to the centralization of the data collection process and the uniformity of sample selection.

Furthermore, the paper uses two approaches to address the problems of quality and unobserved heterogeneity, which plague the literature on university efficiency. First, it exploits the availability of quality rankings to calculate Spearman correlations between predicted efficiency scores and quality measures based on rankings of the QS (2010). Secondly, the paper uses the true random effects stochastic frontier approach proposed by Greene (2005a,b), which exploits the panel data structure to disentangle unobserved heterogeneity and efficiency.

The paper is structured as follows: Section 2 summarizes the existing literature. Sections 3 and 4 describe the data and the applied methodology. Section 5 discusses the estimation results. Section 6 summarizes the paper.

2 Literature

The literature on university efficiency grows rapidly. Worthington (2001) and Johnes (2004) provide literature reviews. However, little evidence in respect to cross-country comparisons exists (Agasisti and Johnes, 2009). Salerno (2003) explains that cross-country comparisons face the difficulties to obtain comparable data and to ensure institutional comparability and sample homogeneity. Hence, only few studies spanning multiple countries exist.

Joumady and Ris (2005) conducted 209 interviews of graduate students across eight European countries and use the resulting information to estimate teaching efficiency of universities using Data Envelopment Analysis (DEA). Aghion et al. (2010) do not estimate a production frontier, but compare the research productivity in US and European universities, showing that the latter lag behind according to a number of indicators. Furthermore, they find that autonomy and accountability boost productivity. Bonaccorsi and Daraio (2007) provide an in-depth analysis of university specialization and performance by exploiting the Aquameth database which contains micro-level information about universities across Europe.

In addition, the existing literature contains a small number of papers providing pairwise country comparisons. Namely, Agasisti and Johnes (2009) employ the DEA methodology to UK and Italian administrative data and demonstrate that technical efficiency of UK universities is higher. Similarly, Agasisti and Pérez-Esparrells (2009) compare the efficiency of Spanish and Italian universities and find higher efficiencies for Italian universities.

3 Data

Based on data from the QS World University Rankings (QS, 2010), this paper estimates an input distance function. Inputs enter as the number of full-time equivalent (FTE) staff.¹ The assumed production technology considers three outputs, namely

¹For observations that only entail information about headcount, the measures for FTE staff and students refer to headcount multiplied by the average ratio between headcount and FTE staff/students.

FTE undergraduate students, FTE graduate students and citations. Citations refer to the score of the ranking item "citations per employee" multiplied by the number of FTE employees.² The employed variables are normalized by the sample median. Table 1 provides summary statistics of the variables.

Table	1:	Summary	statistics
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Variable	Description	Mean	Std. Dev.	Min	Max
Staff	Employees (FTE)	1810.567	1077.974	88	6637
Ugrad	Undergraduate Students (FTE)	16022.98	12367.46	173	130227
Grad	Graduate Students (FTE)	5702.221	3860.372	372.3154	32283
Cit	Index of citations within 5 years	128953.9	83184.15	4576	458874

The sample consists of universities for which an individual score for the item "citations per FTE employee" exists, i.e. the top 300 research universities. Restraining the sample to universities observed in multiple time periods and dropping observations with missing values yields a sample of 273 universities over time, resulting in 720 observations.

4 Methodology

This paper employs two alternative methodologies to identify university efficiencies. The first methodology consists of estimating a fixed effect estimator (*FE*). Schmidt and Sickles (1984) suggest to transform the predicted individual intercepts $(\hat{\alpha}_i)$ by subtracting them from the maximum intercept $(max(\hat{\alpha}_i))$ and to interpret the resulting deviations as inefficiency. This approach has the advantage that it is distribution free except for the normally distributed error term. However, it might suffer from the incidental parameter problem (Lancaster, 2000) and assumes that unobserved heterogeneity comprises only efficiency. A translog specification of the input distance function approximates production technology as specified in formula 1:

$$-\ln x_{it} = \sum_{m=1}^{3} \gamma_m \ln y_{mit} + \frac{1}{2} \sum_{m=1}^{3} \sum_{n=1}^{3} \gamma_{mn} \ln y_{mit} \ln y_{nit} + \delta_t + \alpha_i + \nu_{it}, \qquad (1)$$

The dependent variable (x_{it}) captures the number of FTE employees of university i, at time t. In line with the literature on input distance functions, x_{it} enters with a negative sign. The vector of explanatory variables entails three outputs (y_{mit}) ,

 $^{^2\}mathrm{Hence},$ this measure assumes that the citations per employee of the best university remains constant over time.

namely FTE undergraduate students, FTE graduate students and citations. Year dummies (δ_t) account for differences across time. ν_{it} refers to the traditional error term, i.e. follows a normal distribution with mean zero and variance σ_{ν}^2 . α_i denotes individual intercepts, i.e. university-specific dummy variables. Calculation of predicted technical efficiencies (\widehat{TE}_i) follows

$$\widehat{TE}_i = exp(-\widehat{v}_i) = exp(-(max(\widehat{\alpha}_i) - \widehat{\alpha}_i))$$
(2)

The second methodology to identify efficiency builds on the idea of Aigner et al. (1977) and Meeusen and van den Broeck (1977). The stochastic frontier analysis (SFA) identifies inefficiency by assuming that it follows a half-normal distribution. Using the same production technology as above yields the following econometric specification:

$$-\ln x_{it} = \sum_{m=1}^{3} \gamma_m \ln y_{mit} + \frac{1}{2} \sum_{m=1}^{3} \sum_{n=1}^{3} \gamma_{mn} \ln y_{mit} \ln y_{nit} + \delta_t + \varepsilon_{it}$$
(3)

The error term consists of two parts, i.e. $\varepsilon_{it} = \nu_{it} + v_i$. As before, ν_{it} refers to a normally distributed error term with mean zero and variance σ_{ν}^2 . v_i denotes the time-invariant, half-normally distributed inefficiency term, i.e. $v_i = |U_i|$, with $U_i \sim N(0, \sigma_v)$. The methodology developed by Jondrow et al. (1982) backs out inefficiency scores according to

$$E[\upsilon|\varepsilon] = \frac{\sigma\lambda}{1+\lambda^2} \left[\frac{\phi(z)}{1-\Phi(z)} - z\right], z = \frac{\varepsilon\lambda}{\sigma}$$
(4)

where $\lambda = \frac{\sigma_v}{\sigma_v}$ and $\sigma = \sqrt{\sigma_v^2 + \sigma_v^2}$. As above, the exponential of negative inefficiencies yields technical efficiency scores, i.e. $\widehat{TE}_i = exp(-\widehat{v}_i)$

While the data accounts for the quality of research using citations, a number of potential reasons for unobserved heterogeneity exists, e.g. the quality of students, education and staff. The inability of the data to identify the relative relevance of scientific fields poses an additional problem, since both the adequacy of citations to measure research output as well as the average costs of education differ across fields. Finally, various cross-country differences might influence the estimates.

In order to analyze the relevance of unobserved heterogeneity for the identification of university efficiency rankings, this paper further presents a true random effects stochastic frontier approach (*True RE SFA*), which tackles the problem of unobserved heterogeneity by adding a set of time-invariant, university-specific intercepts, α_i , to formula 3 (Greene, 2005a,b). As the estimators' name suggests, the university-specific intercepts presumably follow a normal distribution with mean μ_{α} and variance σ_{α}^2 . Hence, the estimation can be written as:

$$-\ln x_{it} = \sum_{m=1}^{3} \gamma_m \ln y_{mit} + \frac{1}{2} \sum_{m=1}^{3} \sum_{n=1}^{3} \gamma_{mn} \ln y_{mit} \ln y_{nit} + \delta_t + \alpha_i + \varepsilon_{it}$$
(5)

As before, the error term ε_{it} consists of two parts. ν_{it} refers to a normally distributed error term with mean zero and variance σ_{ν}^2 . Furthermore, the *True RE SFA* assumes that the second component, inefficiency, ν_{it} , varies over time. Hence, the identifying distributional assumption concerning the inefficiency term becomes $\nu_{it} = |U_{it}|$, with $U_{it} \sim N(0, \sigma_{\nu})$. Limdep estimates the *True RE SFA* using a simulated maximum likelihood estimator with 100 draws using Halton sequences. In order to facilitate the comparison across models, the discussion centers around timeinvariant efficiency estimates calculated as the average of yearly efficiency scores.

Table 2 summarizes the assumptions underlying the three estimators employed in this paper:

Table 2: Econometric and distributional assumptions of FE, SFA and True RE SFA

	FE	SFA	True RE SFA
Heterogeneity	0	0	$\alpha_i \sim N$
Inefficiency	$max(\alpha_i) - \alpha_i$	$v_i = U_i , U_i \sim N$	$v_{it} = U_{it} , U_{it} \sim N$

5 Results

The coefficient estimates of the econometric estimations appear in the top panel of table 4. The estimates behave well in the sense that the first-order coefficients of outputs have the expected negative sign. Furthermore, the coefficient estimates remain stable across methodologies. The bottom panel of table 4 displays the Spearman correlations between the predicted efficiency scores of alternative estimation techniques. Table 5 in section 7 displays individual university rankings of predicted efficiency scores for each methodology.

Figure 1 plots the predicted efficiencies of the fixed effects (FE) and stochastic frontier (SFA) approaches sorted by the ranking indicated by the FE model for each country. It reveals that the estimated levels of efficiency in the stochastic frontier framework dominate those predicted by the fixed effects model. However, comparing the ordering of the two estimators indicates a high correlation of rankings. The

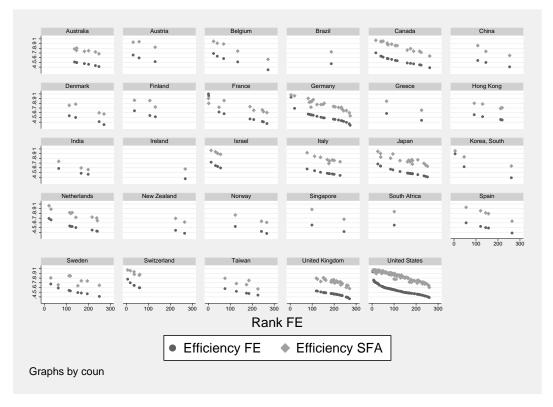


Figure 1: Distribution of efficiency scores across countries comparing FE and SFA

lower panel of table 4 confirms this visual impression by showing a Spearman rank correlation of nearly 0.9 between the FE and SFA model. The high correlation of these two approaches to identify efficiency supports the distributional assumption of the SFA.

In order to facilitate the comparison of predicted efficiencies across countries, table 3 displays the mean and maximum of predicted efficiency scores in each country. However, the interpretation of these indicators requires cautiousness since our sample reflects a particular selection of universities and not the population or a representative drawing thereof.

Both the FE and the SFA estimator suggest that Israel and Switzerland have the highest mean of university efficiency. Furthermore, table 3 displays relatively stable rankings of the ten countries with the highest mean efficiency across methodologies. Calculating the average rank across the two methodologies suggests that Austria ranks third, followed by the USA, South Korea, Finland Canada, South Africa and Belgium. In addition, France and Spain appear within the top ten, but only accord-

Table 3:	Distribution	of	efficiency	estimates	across	countries
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		F.	_	SFA	
Country	N	Mean	Max	Mean	Max
Australia	7	0.47	0.51	0.75	0.80
Austria	3	0.59	0.66	0.90	0.94
Belgium	5	0.55	0.69	0.81	0.95
Brazil	1	0.47	0.47	0.72	0.72
Canada	16	0.54	0.70	0.84	0.97
China	3	0.48	0.54	0.75	0.86
Denmark	4	0.45	0.54	0.68	0.78
Finland	3	0.57	0.64	0.82	0.86
France	9	0.59	1.00	0.73	0.90
Germany	21	0.51	0.93	0.76	0.98
Greece	2	0.51	0.59	0.75	0.85
Hong Kong	5	0.49	0.56	0.74	0.80
India	3	0.51	0.59	0.63	0.74
Ireland	1	0.37	0.37	0.58	0.58
Israel	4	0.65	0.72	0.93	0.97
Italy	9	0.49	0.58	0.78	0.92
Japan	18	0.51	0.68	0.77	0.95
Korea, South	3	0.64	0.90	0.81	0.96
Netherlands	9	0.53	0.69	0.79	0.96
New Zealand	2	0.41	0.44	0.66	0.69
Norway	3	0.44	0.52	0.67	0.77
Singapore	2	0.48	0.55	0.78	0.89
South Africa	1	0.55	0.55	0.84	0.84
Spain	5	0.50	0.60	0.81	0.93
Sweden	8	0.52	0.67	0.74	0.85
Switzerland	6	0.66	0.78	0.91	0.97
Taiwan	5	0.50	0.58	0.69	0.80
United Kingdom	28	0.45	0.53	0.72	0.81
USA	87	0.56	0.93	0.85	0.97
Total	273	0.53	1.00	0.79	0.98

The table shows the number of observations in each country as well as the mean and maximum of predicted efficiencies according to the fixed effects (FE) and stochastic frontier (SFA) methodologies.

ing to one of the two estimators. Comparing the results of medium ranked countries across the two methodologies reveals a relatively volatile order. However, the ranking turns more stable for the bottom third again, where Brazil ranks before the UK, Denmark, Norway, New Zealand and Ireland. The appearance of China, France, Singapore and Taiwan in the bottom third of the distribution depends on the employed methodology.

Comparing these results to those obtained by Journady and Ris (2005) suggests relatively consistent efficiency estimates despite the fact that Journady and Ris (2005) estimate teaching efficiency using the non-parametric data envelopment approach, while this paper employs a parametric framework accounting for both teaching and research. Among the overlapping countries, only two display differences between these two papers. Namely, Finland ranks low in Journady and Ris (2005) but high according to the above results. Conversely, the UK performs well in Journady and Ris (2005) but not in this paper.

The second measure, the maximum efficiency of a university within a country, reflects the minimum distance to the frontier and hence allows to identify the countries which form the production frontier. Similar to the comparison of mean efficiencies across methodologies, the minimum distance to the frontier appears relatively stable among the top ten countries. However, figure 1 reveals that the distribution of efficiencies within a country differs between the two methodologies. Concretely, the maximum efficiency within a country drops relatively quickly for the FE methodology but remains high throughout the top universities according to the SFA estimator. Both estimators locate Germany and USA at the frontier. The average of ranks across methodologies places Switzerland next, followed by Israel, South Korea, Canada, the Netherlands, Belgium and Japan. France displays a high volatility as it achieves the highest value using the FE estimator but merely reaches the 13th rank based on the SFA methodology. The volatility of estimates across methodologies remains high for both the middle and the bottom third of countries.

However, the FE and SFA methodologies do not account for quality and unobserved heterogeneity. Hence, the lower panel of table 4 provides first indication whether unobserved heterogeneity biases the estimation results. It shows information about the Spearman correlations of estimated efficiency scores and measures of quality based on the QS (2010). Concretely, the displayed quality measures refer to the inverse of rankings in terms of peer and employer surveys as well as the inverse of the overall QS quality ranking. As table 4 shows, the correlations are low and even mostly positive, despite the fact that failing to account for quality adequately suggests a negative correlation between efficiency and quality. Hence, these correlations provide indicative evidence that the employed econometric methodologies account for quality in a sufficient manner.³

In order to provide a more thorough analysis of unobserved heterogeneity, table 4 further portrays the coefficient estimates of the *True RE SFA* in column 3. The estimates behave well in the sense that the first-order coefficients of outputs have the expected negative sign and resemble closely those of the *FE* and *SFA* methodologies. However, the estimated standard deviation of efficiency, σ_v , drops to merely 0.00001. Hence, the predicted efficiency scores barely vary, implying a negligible identification power of this estimator. The *True RE SFA* methodology essentially divides universities into four categories. The "Universities. Then, the estimator creates a bulk of universities that cannot be distinguished. Finally, eleven universities constitute the rear.

However, table 4 further reveals that the True RE SFA yields efficiency esti-

³Similarly, including quality measures in the estimation directly supports this interpretation as well. Concretely, the coefficients of quality measures turn out insignificant or even positive. Furthermore, the Spearman correlation of the resulting efficiency estimates with the unadjusted FE and SFA scores remains above 0.95.

mates with a Spearman correlation of more than 0.85 compared to the FE and SFA methodologies. The high Spearman correlation suggests that accounting for unobserved heterogeneity might not be as relevant to obtain credible efficiency ranking estimates. Hence, the robustness check of estimating a *True RE SFA* suggests that interpreting the predictions of the simpler methodologies FE and SFA appears adequate.

An interesting interpretation of the applied estimation methodologies follows Greene (2004), according to which the FE estimator underestimates efficiency since it labels all unobserved heterogeneity as inefficiency. The *True RE SFA* estimator on the other hand tends to capture too much of between variation in unobserved heterogeneity and hence overestimates efficiency. This interpretation suggests that the *FE* and *True RE SFA* estimators form the lower and upper boundaries for the true efficiency scores, respectively. Figure 2 displays the predicted efficiency scores for each of the three estimation techniques sorted by the *FE* efficiency scores. The predictions of the *SFA* lie within the predictions of the *FE* and the *True RE SFA* estimator. Hence, these results support the above interpretation as efficiency boundaries. Thereby, figure 2 provides further indicative evidence for the hypothesis that the *SFA* methodology yields reasonable efficiency predictions.

6 Summary

This paper applies the production function framework to an international perspective of universities by exploiting data provided by the QS World University Rankings (QS, 2010) to estimate a multi-output input distance function for 273 universities in 29 countries between 2007 and 2009. Thereby the paper extends the existing literature by providing first micro-level evidence of the global university production frontier.

The estimated input distance function uses staff to measure inputs. Outputs refer to undergraduate students, graduate students and citations. A translog specification approximates the production technology. The paper contrasts two strategies to identify technical efficiency. First, the deterministic frontier approach proposed by Schmidt and Sickles (1984) estimates a fixed effect estimator. Assuming that the highest predicted individual intercept reflects the most efficient university allows to interpret transformed intercepts as inefficiency. Secondly, the stochastic frontier approach identifies technical efficiency by assuming that it follows a halfnormal distribution (Aigner et al., 1977; Meeusen and van den Broeck, 1977). The two methodologies yield relatively similar efficiency rankings estimates.

The predicted efficiency scores reveal that Israel and Switzerland display the highest average efficiency, followed by Austria, the USA, South Korea and Finland.

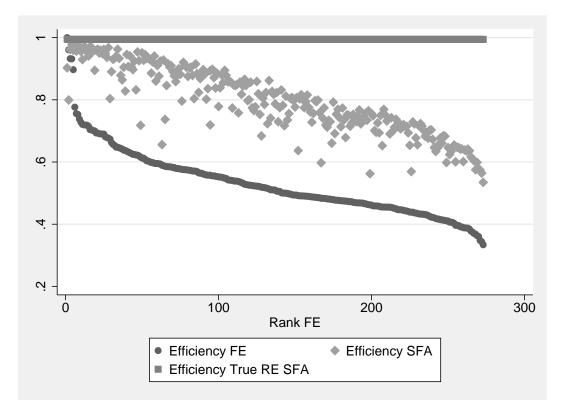


Figure 2: Estimation boundaries of efficiency scores

Germany and the USA, pursued by Switzerland, Israel and South Korea, form the production frontier.

Furthermore, the paper uses two approaches to address the problems of quality and unobserved heterogeneity which plague the literature on university efficiency. First, the finding that the Spearman correlation between technical efficiency scores and quality measures based on rankings of the QS (2010) is low or even positive suggests that the fixed effects and stochastic frontier approach account for quality in a sufficient manner. Secondly, the paper uses the true random effects stochastic frontier approach proposed by Greene (2005a,b), thereby exploits the panel data structure to disentangle unobserved heterogeneity and efficiency. The Spearman correlations between the three employed estimators remain high after accounting for unobserved heterogeneity, suggesting that simple estimation techniques suffice to obtain credible efficiency ranking estimates. However, by revealing that the true random effects stochastic frontier yields statistically uninformative efficiency estimates, this econometric exercise also illustrates the challenges in providing adequate information to university managers to asses universities.

1

7 Tables

Table 5: Ranks of predicted university efficiencies according to FE, SFA and True RESFA

University	;	FE	SFA	TRE	University	1	FE	SFA	TRE
AUSTRALIAN National Uni	J AU	146	162	147	University CHALMERS University of Tech	J	7 E 29	123	147
University of ADELAIDE, The	AU	182	182	147	Kth, ROYAL INSTITUTE of Tech	SE	152	253	147
University of MELBOURNE	AU	232	195	147	LUND University	SE	118	99	147
University of NEW SOUTH WALES	AU	147	122	147	STOCKHOLM School of Economics	SE	63	238	147
University of QUEENSLAND,	AU	201	170	147	STOCKHOLM University	SE	195	179	147
University of SYDNEY	AU	248	228	147	UPPSALA University	SE	112	97	147
Uni of WESTERN AUSTRALIA	AU	137	138	147	Umeå University	SE	168	174	147
MCI Management Center INNSBRUCK	AT	31	37	17	Uni of GOTHENBURG	SE	250	245	147
University of GRAZ	AT	131	111	147	ETH Zurich	CH	35	35	147
University of VIENNA	AT	58	30	147	ETH LAUSANNE	CH	36	76	147
Catholic University of LEUVEN	BE	70	67	147	University of BERN	CH	61	81	147
Free University of Brussels(VUB)	BE	134	175	147	University of GENEVA	CH	18	18	17
University of GHENT	BE	272	271	267	University of LAUSANNE	CH	59	88	147
University of LIEGE	BE	24	25	17	University of ZURICH	CH	6	5	17
Universite Catholique de LOUVAIN	BE	41	52	147	NATIONAL TAIWAN Uni	TW	185	168	147
State University of CAMPINAS	BR	187	194	147	National CHIAO TUNG Uni	TW	174	235	147
DALHOUSIE University	CA	71	80	147	National SUN YAT-SEN Uni	TW	128	227	147
LAVAL University	CA	89	69	147	National TSING HUA Uni	TW	76	134	147
Mcgill University	CA	262	255	147	National YANG MING Uni	TW	226	270	267
Mcmaster University	CA	17	8	17	CARDIFF University	UK	259	249	147
QUEEN'S University	CA	160	156	147	DURHAM University	UK	162	160	147
SIMON FRASER University	CA	189	177	147	IMPERIAL College London	UK	159	144	147
The University of WESTERN ONTARIO	CA	40	26	17	KING'S College London	UK	222	209	147
University of ALBERTA	CA	216	180	147	NEWCASTLE University	UK	237	233	147
University of BRITISH COLUMBIA	CA	175	137	147	Queen's Uni of BELFAST	UK	270	267	267
University of CALGARY	CA	42	28	147	UCL	UK	153	125	147
University of MANITOBA	CA	223	208	147	University of ABERDEEN	UK	202	218	147
University of OTTAWA	CA	114	94	147	University of BATH	UK	203	220	147
University of TORONTO	CA	50	22	147	University of BIRMINGHAM	UK	196	159	147
University of VICTORIA	CA	73	79	147	University of BRISTOL	UK	156	146	147
University of WATERLOO	CA	69	51	147	University of CAMBRIDGE	UK	151	120	147
Université De Montréal	CA	106	92	147	University of DUNDEE	UK	193	197	147
FUDAN University	CN	253	243	147	University of EAST ANGLIA	UK	255	263	267
SHANDONG University	CN	108	93	147	University of EDINBURGH	UK	211	176	147
University of S&T of China	CN	145	183	147	University of GLASGOW	UK	188	165	147
AARHUS University	DK	142	140	147	University of LEEDS	UK	257	247	147
Technical Uni of DENMARK	DK	113	158	147	University of LEICESTER	UK	119	131	147
Uni of COPENHAGEN	DK	271	269	147	University of LIVERPOOL	UK	215	212	147
Uni of SOUTHERN DENMARK	DK	249	265	147	University of MANCHESTER	UK	228	199	147
KUOPIO University	FI	37	86	147	University of NOTTINGHAM	UK	263	251	267
University of HELSINKI	FI	107	90	147	University of OXFORD	UK	198	155	147
University of TURKU	FI	132	196	147	University of READING	UK	230	239	147
Claude Bernard University Lyon 1	FR	190	187	147	University of SHEFFIELD	UK	234	217	147
Ecole Normale Superieure de LYON	FR	2	129	147	University of SOUTHAMPTON	UK	210	186	147
Ecole Normale Superieure, Paris	FR	1	59	147	University of ST ANDREWS	UK	139	189	147
Joseph Fourier Uni - GRENOBLE I	FR	251	258	267	University of SUSSEX	UK	124	142	147
Paris-Sud XI University	FR	246	236	147	University of YORK	UK	186	207	147
Polytechnic School (France)	FR	49	203	147	ARIZONA STATE University	USA	206	169	147
University MONTPELLIER II	FR	207	242	147	BOSTON University	USA	135	109	147
Uni Pierre and Marie Curie	FR	267	261	267	BRANDEIS University	USA	19	65	17
University of Strasbourg	FR	72	91	147	BROWN University	USA	14	21	17
BIELEFELD University	DE	140	145	147	CARNEGIE MELLON Uni	USA	44	63	147
Free University of BERLIN	DE	82	55	147	CASE WESTERN RESERVE Uni	USA	109	149	147
University of Erlangen-Nuernberg	DE	150	143	147	COLORADO STATE University	USA	57	45	147
University of JENA	DE	243	248	147	COLUMBIA University	USA	90	62	147
Goethe Uni FRANKFURT	DE	197	184	147	CORNELL University	USA	62	43	147
HEIDELBERG University	DE	220	213	147	Caltech	USA	3	13	17
Heinrich Heine Uni of Dusseldorf	DE	98	104	147	DARTMOUTH College	USA	15	29	17
Johannes Gutenberg Uni of MAINZ	DE	20	16	17	DREXEL University	USA	260	262	147
Ludwig Maximilian - Uni of MUNICH	DE	155	127	147	DUKE University	USA	86	72	147
Ruhr University BOCHUM	DE	256	244	147	EMORY University	USA	75	74	147
Saarland University	DE	93	105	147	FLORIDA STATE University	USA	213	173	147

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NORWEGIAN University of S&T NO 242 256 147 University of Wisconsin-Madison USA 52 27	147
University of BERGEN NO 123 154 147 VANDERBILT University USA 229 231	
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National University of Singapore SG 99 71 147 WASHINGTON Uni In St. Louis USA 45 44	17
University of CAPE TOWN ZA 101 103 147 YALE University USA 88 84	147
Autonomous Uni of BARCELONA ES 264 254 147	
Autonomous Uni of MADRID ES 55 38 147	
Uni De Santiago De Compostela ES 144 118 147	
University of BARCELONA ES 121 95 147	
University of VALENCIA ES 157 135 147	
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Estimation Results									
depvar: -Staff	FE	SFA	True RE SFA						
Ugrad	-0.049	-0.1062***	-0.096***						
- 8	(0.040)	0.0141	0.0072						
Grad	-0.034	-0.0342***	-0.0363***						
Grad	(0.029)	0.0124	0.0082						
Cit	-0.821***	-0.8537***	-0.8333***						
	(0.033)	0.0174	0.0088						
Ugrad_Ugrad	-0.040**	-0.0317***	-0.0244***						
- 8	(0.016)	0.0094	0.0049						
Grad_Grad	-0.052	-0.0302**	-0.0408***						
Gradiorad	(0.033)	0.0147	0.0077						
Cit_Cit	0.002	0.0057	0.0134						
0101010	(0.031)	0.0143	0.0096						
Ugrad_Grad	0.018	0.0163	0.0364***						
ogradiedrad	(0.032)	0.0177	0.0107						
Ugrad_Cit	-0.004	0.0295*	0.0036						
egrad_ent	(0.034)	0.0167	0.0105						
Grad_Cit	0.055	0.0133	0.0215						
Gradient	(0.046)	0.023	0.0143						
Year_2	-0.151***	-0.1473***	-0.1473***						
i cai <u>1</u> 2	(0.012)	0.0149	0.0115						
Year_3	-0.152***	-0.1516***	-0.1528***						
Teal_3	(0.012)	0.0135	0.0103						
Constant	0.114	0.371***	0.0105						
Constant	(0.071)	0.0186							
Lambda	(0.071)	2.3842	0.0006						
Sigma(v)		0.2866	0.00001						
$\operatorname{Sigma}(\nu)$		0.2800	0.1167						
$\mu(\alpha_i)$		0.1202	0.1296						
			0.1296 0.1574						
$\sigma(\alpha_i)$	720	720	720						
	. = .	Yes							
SFA	No	Yes	Yes Yes						
University Effects	Yes	No	res						
	a								
Spearman Correlations									
	FE	SFA	True RE SFA						
FE	1								
SFA	0.9165	1							
True RE SFA	0.8519	0.8982	1						
Peer Review	0.1043	0.2403^{**}	0.1552^{**}						
Employer Review	-0.0178	0.022	-0.0409						

 Table 4: Estimation results and Spearman correlations
 Particular