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Measuring Individual Risk Attitudes in the Lab: Task or Ask? An Empirical Comparison

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Measuring Individual Risk Attitudes in the Lab: Task or Ask? An Empirical Comparison

Jan-Erik Lönnqvist
Markku Verkasalo
Gari Walkowitz
Philipp C. Wichardt

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Measuring Individual Risk Attitudes in the Lab: 
Task or Ask? An Empirical Comparison

Abstract

This paper compares two prominent empirical measures of individual risk attitudes - the Holt and Laury (2002) lottery-choice task and the multi-item questionnaire advocated by Dohmen, Falk, Huffman, Schupp, Sunde and Wagner (2011) - with respect to (a) their within-subject stability over time (one year) and (b) their correlation with actual risk-taking behaviour in the lab - here the amount sent in a trust game (Berg, Dickaut, McCabe, 1995). As it turns out, the measures themselves are uncorrelated (both times) and, most importantly, only the questionnaire measure exhibits test-re-test stability ($p = .78$), while virtually no such stability is found in the lottery-choice task. In addition, only the questionnaire measure shows the expected correlations with a Big Five personality measure and is correlated with actual risk-taking behaviour. The results suggest that the questionnaire is the more adequate measure of individual risk attitudes for the analysis of behaviour in economic (lab) experiments. Moreover, with respect to behaviour in the trust game, we find a high re-test stability of transfers ($p = .70$). This further supports the conjecture that trusting behaviour indeed has a component which itself is a stable individual characteristic (Glaeser, Laibson, Scheinkman and Soutter, 2000).

JEL-Code: D810, C910, Z100.

Keywords: risk attitudes, trust, personality, lab experiments.

Jan-Erik Lönnqvist  
University of Helsinki / Finland  
jan-erik.lonnqvist@helsinki.fi

Markku Verkasalo  
University of Helsinki / Finland  
markku.verkasalo@helsinki.fi

Gari Walkowitz  
University of Cologne / Germany  
gari.walkowitz@uni-koeln.de

Philipp C. Wichardi*  
Chair of Microeconomics  
University of Rostock  
Ulmenstr. 69  
Germany – 18057 Rostock  
philipp.wichardi@uni-rostock.de

*corresponding author

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1 Introduction

Risk-aversion has long since been a standard ingredient of economic theory (seminal 
works being Arrow, 1965, and Pratt, 1964). Moreover, with the growing body of 
experimental studies in the social sciences in general and in economics in particular 
(see Falk and Heckman, 2009, for a discussion), individual risk attitudes have also 
been linked empirically to a good many behavioural patterns. For example, Goeree, 
Holt and Palfrey show how the distribution of bids in a private-values auction experi-
ment, including the prevalence of overbidding, can be captured in a quantal response 
equilibrium model with risk averse bidders (Goeree et al., 2002), and how deviations 
from Nash-equilibrium in a generalised matching pennies game can be explained if 
the subjects’ risk aversion is taken into account (Goeree et al., 2003). Dohmen and 
Falk (2011) in turn find that individuals self-select into different types of payment 
schemes according to their risk attitudes with more risk averse individuals choosing 
less performance dependent payment schemes (see also Guiso and Paiella, 2005). And 
these are but some of the many examples that have been discussed in the literature 
(see Harrison and Rutström, 2008, for further examples).

Despite the empirical relevance of individual risk attitudes, however, there is still 
an ongoing debate about their adequate elicitation (cf. Harrison and Rutström, 
2008). Among other aspects — including specific procedural issues as well as the 
general question whether risk aversion is at all a plausible phenomenon for the small 
stakes which are commonly offered in the lab (cf. Rabin, 2000) — the relevance of 
proper incentives has given rise to discussions: Should risk preferences be inferred 
from incentivised behavioural measures such as lottery choice tasks (e.g. Holt and 
Laury, 2002) as commonly done in economics? Or should they rather be assessed 
using non-incentivised questionnaires based on so called Likert statements in which 
subjects specify their level of agreement to a certain statement as predominantly used 
in psychology (see Lauriola and Levin, 2001, for a historical review; see also Dohmen, 
Falk, Huffman, Schupp, Sunde and Wagner, 2011)?

Of course, lottery choices are easier to translate into formal indices and, hence, are 
preferable when questions about structural parameters of utility functions are at issue 
(see, for example, Andersen, Harrison, Lau and Rutström, 2008a). Over the years, 
however, a mounting evidence has gathered indicating that these measures come with 
a considerable degree of noise, which may vary with the exact method of elicitation 
(e.g. Isaac and James, 2000; Berg, Dickhaut and McCabe, 2005; Hey, Morone and 
Schmidt, 2009, Dulleck, Fell and Fookem, 2011) and/or the subjects cognitive ability
As a result of this, also questionnaire measures have witnessed a growing popularity in recent years (e.g. Dohmen and Falk, 2011; Dohmen, Falk, Huffman and Sunde, 2008 and 2010). While inherently difficult to translate into numerical parameters, they still have their virtues as predictors of behaviour. This is in particular so for the analysis of behaviour in lab experiments, where relative assessments of risk attitudes are often more important than absolute ones. And, of course, questionnaire measures are both cheaper to use and arguably easier to respond to without further instructions. This notwithstanding, the correlation between experimental and questionnaire / survey measures again tends do be weaker than what one might have hoped for (see, for example, Anderson and Mellor, 2009) – thereby also raising questions about the stability of the preferences thus expressed.

In the present study, we take up the question about the adequate elicitation of risk attitudes in connection with economic lab experiments and investigate the relative performance of a lottery-choice measure validated by Holt and Laury (2002) and a questionnaire measure advocated by Dohmen et al. (2011) with regard to the correlation between them, their construct validity (being correlated with an external predictor of risk-taking behaviour, namely personality), and their test-re-test stability over time (one year). Moreover, we test the ability of the two types of measures to predict actual behaviour in a domain that is typically associated with risk.1 In doing so, we operationalise risk-taking behaviour in the lab via a standard trust / investment game (Berg, Dickhaut and McCabe, 1995) taking the amount sent by first movers as a measure for their willingness to take risks. Although the question whether the investment decision in trust games eventually measures risk or trust (or both) is contentious — with some studies emphasising the role of risk (e.g. Snijders and Keren, 1998, Ben-Ner and Putterman, 2001) and others rather questioning it (e.g. Eckel and Wilson, 2004, or Bohnet and Zeckhauser, 2004) or arguing directly in favour of trust (e.g. Houser, Schunk and Winter, 2010) — it appeared to us as an intuitive item to analyse. It fact, the popular view that giving money to strangers is risky (also expressed, for example, in Kosfeld, Heinrichs, Zak, Fischbacher and Fehr,

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1Note that biases in the measurement – as might result from one of two measures is incentivised while the other is not – are not problematic for our purposes as we “only” intend to ascertain which of the two (well established) measures is most suited as a tool to relate observed behaviour in economic (lab) experiments to individual risk attitudes but do not draw any inferences about the shape of individual utility functions. For an interesting discussion of hypothetical biases in the elicitation of risk attitudes and a review of the respective literature, see Harrison (2006).
2005) intuitively suggests that doing so should be correlated with risk attitudes. In conjunction with the mixed evidence on this correlation, it thus seemed appealing to try and scrutinise whether the type of risk measure applied may be crucial in this context.

In the experiment we conducted, subjects first played a standard one shot trust game – the main experimental task. Once that was completed, subjects were given the Holt and Laury (2002) lottery choice task and, eventually, the questions about individual risk-attitudes advanced by Dohmen et al. (2011) as part of the post experimental questionnaire. Moreover, in order to further analyse the connection between risk attitudes and other fundamental personal characteristics, we also gathered data on the subjects’ Big Five personality traits (Neuroticism, Extraversion, Openness, Agreeableness, Conscientiousness), which subsume a huge variety of personality attributes and provide a concise summary of stable individual differences in personality (Digman, 1990). These traits had previously been shown to be correlated with risk-taking behaviour (Nicholson, Soane, Fenton-O’Creery and Willman, 2005; or Lauriola and Levin, 2001), so that it seemed plausible to expect them to be correlated with the measures of individual risk attitudes under scrutiny, i.e. with behaviour in the lottery-choice task and the answers to the questionnaire. The data on personality where gathered prior to and independent of the actual (first) experiment, though, in order to obtain an additional external reference point against which to compare our later results. Finally, in order to further check on the stability of our results, we reran the actual experiment with (some of) the same subjects after about a year’s time and compared test and re-test results.

The results show the following general patterns: (1) The two measures of risk-attitudes themselves are uncorrelated (both studies). (2) Both the questionnaire measure (all single items and a factor that was statistically extracted from the questionnaire measure about risk attitudes) and individual risk-taking (trusting) behaviour exhibit a high and significant test-re-test stability, while virtually no such stability is

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2For an instructive general discussion of the different facets of trust see, for example, Ben-Ner and Halldorsson (2010).

3Risk-measures were gathered after the trust game (the main behavioural task) and without being explicitly announced before in order to avoid any external indications that trusting behaviour is supposed to be risky. It is noteworthy in this respect that Houser et al. (2010), who ran a similar experiment combining a trust game with the Holt and Laury lottery-choice task in changing order, do not find any order effects. Inasmuch as possible, we also do not find any evidence in this direction in our data (see Section 3).

4The Big Five personality measure is also used, for example, by Bartling, Fehr, Marechal and Schunk (2005) who provide evidence that Agreeableness is related to a person’s competitiveness.
present in the lottery-choice task. (3) The factor derived from the questionnaire as well as all single items are correlated with the Big Five personality measure, while behaviour in the lottery-choice task is mostly not. (4) The investment decision in the trust game is correlated with the risk factor as well as with some of the single items, but is virtually uncorrelated with the subjects’ choices in the Holt and Laury lottery-choice task (this holds for both test and re-test); and we do not find any evidence for hedging or reference point effects that could explain the lack of correlation with the subjects’ lottery choices (see Section 3 for details).

From these observations, we conclude that the type of multi-item questionnaire advocated by Dohmen et al. (2011) is a useful and reliable tool to measure individual risk attitudes in connection with economic lab experiments, whereas the behavioural measure, i.e. the Holt and Laury (2002) lottery-choice task with small stakes, appears to be comparably unreliable in this respect.\footnote{While most lab-experiments indeed rely on small stakes when using lottery choice tasks to study subjects characteristics, evidence for high-stakes lottery choices of 31 subjects followed over 6 months indicates that this may be more stable (Harrison, Johnson, McInnes and Ruström, 2005). Whether the difference in stability is due to the higher stakes, the shorter period between test and re-test, unobserved differences in the subject pool or simply chance is difficult to judge, though.}

Moreover, regarding the discussion about the determinants of trust (as measured in the Berg-Dickhaut-McCabe trust game), we interpret the data as suggesting that individual risk attitudes, once elicited in the appropriate way, indeed play a significant role in such behaviour (in particular attitudes towards “risk in financial matters,” see Section 3 for details). Last but not least, we see the high over-time stability of investment behaviour in the trust game as a strong indication that trusting behaviour indeed has a component which is a stable individual characteristic as conjectured by Glaeser, Laibson, Scheinkman and Soutter (2000, p. 827).\footnote{In fact, this observation also squares well with the results by Cesarini et al. (2008) which indicate a correlation between trusting behaviour and the subjects genetic code; see also Fehr (2009) for a recent summary on the literature on the determinants of trust.}

The remainder of this paper is structured as follows: In Section 2, we present the experimental design and procedures. The empirical results of our study are gathered and briefly discussed in Section 3. Section 4 concludes.

2 Experimental Design and Procedures

In the sequel, we describe the design of the lab-experiment (2.1) and the procedures of the overall study (2.2).
2.1 Design of the Experiment

The lab-experiment consisted of three different parts: The Holt and Laury lottery-choice task, a post experimental questionnaire including the questionnaire measure of risk attitudes, and the trust game played at the beginning of the experimental sessions. These are described below.

The Lottery-Choice Task
As an incentivised behavioural measure of individual risk attitudes, we used the well established lottery choice task proposed by Holt and Laury (2002). This measure presents participants with ten choices between paired lotteries with payoffs ranging from .10 to 3.85 Euros (see Table 1 in Holt and Laury, 2002, P. 1645, who pay in US-Dollars instead of Euros). The first four pairs of lotteries give a higher expected payoff for the safer choice, whereas the latter six pairs give a higher expected payoff for the more risky choice. Thus, the risk neutral option is to make four safe choices and then six risky choices. At the end of the experiment, one of the ten pairs of lotteries is randomly selected, the preferred lottery of that pair is conducted and the subject is payed accordingly. Following Holt and Laury (2002), we took the number of risky decisions the participant made as our measure of individual risk attitudes.

Asking About Risk Attitudes - the questionnaire measure
As a non-incentivised questionnaire measure of individual risk-attitudes, we adhered to the questionnaire items presented by Dohmen et al. (2011) and asked subjects to express their willingness to take risks in (1) general (2) driving (3) financial matters (4) sport and leisure activities (5) career decisions (6) health behaviours and (7) trusting strangers, using an eleven-point scale with zero indicating complete unwillingness to take risks and ten indicating complete willingness to take risks. This type of questionnaire measure has been shown by Dohmen et al. to predict risky behaviors in various life domains. As our measure of individual risk attitudes, we used a factor which was statistically extracted from the answers to all seven questions (using a

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7See also Harrison, Johnson, McInnes and Rutström (2005a), and Holt and Laury (2005).
8We also conducted the subsequent analyses using the point at which subjects crossed over from the safe choice to the risky choice as our measure of risk-taking (215 subjects, i.e. 93%, had such a crossover point and did not switch back from the risky choices to the safe choices). However, as the results were virtually identical, we report only the results for the former type of coding.
9The last item is not part of the questions initially developed for the German Socio-Economic-Panel (SOEP) and analysed by Dohmen et al. (2011) but was added for purposes of the present study. In view of the later discussion, it is noteworthy that its inclusion into the factor model is not crucial for the results.
standard factor analysis), henceforth referred to as the general risk factor - GRF for short.\textsuperscript{10} However, in the sequel, we occasionally report also correlations of single items.

The Trust Game – Risk-associated behaviour

The main experimental task for the subjects was the type of trust game first proposed by Berg, Dickhaut and McCabe (1995). In this game, there are two players, an investor and a responder. Initially, all players were given 10 Talers (1 Taler = .60 Euro). Responders just pocketed the money. By contrast, investors were given the opportunity to transfer any non-negative integer part of the endowment to a randomly matched and anonymous responder. All transfers were tripled by the experimenter. Upon receiving the transfer, responders had to decide how much to return to the investor; responders’ back-transfers, which were elicited with the strategy method (Selten, 1967), i.e. without any information about actual transfers, are not analysed in the present paper. The amount transferred by investors was taken as our measure of risk-taking behaviour. The trust game concluded with the elicitation of senders’ and responder’s first order beliefs which are partly reported in the sequel.

2.2 Procedures

The study was conducted in three phases. In the first phase, participants’ personality profiles were gathered using a questionnaire posted on a University of Helsinki website;\textsuperscript{11} the questionnaire also asked about some further items which are unrelated to the present study as well as some statistical data which allowed us to track subjects over time. The questionnaire was open from November 15\textsuperscript{th} to December 6\textsuperscript{th} 2007. Participants were contacted through ORSEE (Greiner, 2004) based on a mailing list of the Laboratory for Experimental Economics (University of Bonn) to which they had signed up in order to take part in research conducted at the lab. The mailing list consisted of around 3000 names. As an incentive to participate, participants were offered feedback on their personality profiles and a potential payoff from a short non-related decision task placed after the questionnaires; 945 subjects participated in this phase.

\textsuperscript{10}Parallel analysis revealed that the questionnaire data can be modelled by just one principal component that accounted for 39% of the variance. In the present setting, results for the factor, which is a kind of (normalised) sophisticated average, are almost identical to those for the standard average, though.

\textsuperscript{11}Personality was measured with a 60-item questionnaire (Konstabel, Lönnqvist, Walkowitz, Konstabel and Verkasalo, 2012) that captures the five personality factors conceptualised in the Five-Factor Model (Costa and McCrae, 1992).
Out of the 945 subjects for which personality profiles had been gathered, a total of 232 responded to our subsequent invitation (which made no reference to phase 1) and participated in the experimental sessions conducted in phase 2 of our study at the Laboratory for Experimental Economics in Bonn between December 2007 and February 2008 (109 male, mean age 23.5 years, $SD$ of age $= 3.2$). Upon arriving at the lab, subjects were seated in computer cubicles and, thereby, randomly assigned to either of two roles for the trust game so that we ended up with $N = 116$ investors and equally many responders. The experiment, which was implemented using z-tree (Fischbacher, 2007), started once the subjects had read the instructions and (correctly) answered a set of control-questions. After decisions for the one shot trust game had been made, all subjects had to complete the Holt and Laury lottery-choice task and, eventually, the questionnaire measure about risk attitudes as part of the post experimental questionnaire (i.e. both measures had not been announced before as they were not part of the primary experiment). Feedback about the actual outcome of the trust game and the lottery was given individually at the end of the experiment.

In phase 3, we reinvited subjects from the 232 participants of phase 2 for a new experiment which took place on March 16 2009: 44 out of the 232 subjects from phase 2 participated in phase 3 (22 male, mean age 24.0 years, $SD$ of age $= 4.1$). The actual experiment was identical to the one described in phase 2 except for the fact that the assignment of subjects to their roles in the trust game was not random but determined by phase-2-assignments. At the very last, subjects were asked whether they had taken part in a similar experiment before (without suggesting that they had).

3 Results

In the sequel, we report the results of the our experiments from phase 2 (the test) and phase 3 (the re-test) of our study focusing in particular on the correlation between the different measures of individual risk attitudes, their correlation with an external predictor of risk-taking behaviour and with investment decisions in the trust game. The results are collapsed across genders as we do not find gender to be correlated with

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12To be precise, 252 subjects participated in the experiments. However, due to technical problems, the first session happened to end up without the questionnaire measure of individual risk attitudes. The session was excluded for the later analysis (this also affects the re-test to which 2 additional subjects from the respective session turned up but were excluded from the analysis).

13The additional items were added at the very end of the post-experimental questionnaire in order to avoid any interference with the earlier aspects of the experiment.
either the lottery-choice behaviour, the general risk factor\textsuperscript{14} or trusting behaviour.

\subsection{The Test Study}

As a first result, we find that the lottery-choice measure and the GRF themselves are uncorrelated (Spearman’s $\rho = -.04$, $p > .10$)\textsuperscript{15} suggesting that they measure different things; see Figure 1 for illustration.

![Figure 1: Decisions in the lottery-choice task (horizontal axis) and the risk questionnaire factor score (GRF; vertical axis). OLS regression line (beta = -.0001) fitted to the data.](image)

In order to find out more about the reliability of the respective risk-attitude measures, we analysed the correlations with the Big Five personality factors; see Table 1. Based on a self-report measure of risk-attitudes in six decision domains Nicholson, Soane, Fenton-O’Creevy and Willman (2005) find that risk-taking is associated positively with Extraversion and Openness, and negatively with Neuroticism, Agree-

\textsuperscript{14}Gender was correlated with one single item of the underlying questionnaire, though, namely the willingness to take risks in financial matters. For this the average for females was 2.642 as opposed to 3.553 for males (a Mann Whitney Test shows that those groups are different, $p < .01$)

\textsuperscript{15}All reported correlations are Spearman rank correlation coefficients of two-sided tests; if not separately specified, Pearson correlation coefficients are almost identical.
ableness and Conscientiousness (see also Lauriola and Levin, 2001). Accordingly, we expected similar relations to hold for the risk-attitude measures tested in the present study. Yet, as shown in Table 1, all five personality factors significantly predict the GRF but do not predict behaviour in the lottery-choice task in multiple regression analyses which model the respective risk-attitude as a composite of the Big Five. Thus, these findings cast some doubt on the reliability of the lottery-choice task as a measure of individual risk attitudes.

<table>
<thead>
<tr>
<th>Measure</th>
<th>(1) LM</th>
<th>(2) GRF</th>
</tr>
</thead>
<tbody>
<tr>
<td>Neuroticism</td>
<td>-.0249</td>
<td>-.0672***</td>
</tr>
<tr>
<td></td>
<td>(.0247)</td>
<td>(.0223)</td>
</tr>
<tr>
<td>Extraversion</td>
<td>.0333</td>
<td>.0500***</td>
</tr>
<tr>
<td></td>
<td>(.0208)</td>
<td>(.0186)</td>
</tr>
<tr>
<td>Openness</td>
<td>-.0104</td>
<td>.0801***</td>
</tr>
<tr>
<td></td>
<td>(.0228)</td>
<td>(.0205)</td>
</tr>
<tr>
<td>Agreeableness</td>
<td>.0323</td>
<td>-.0515**</td>
</tr>
<tr>
<td></td>
<td>(.0248)</td>
<td>(.0223)</td>
</tr>
<tr>
<td>Conscientiousness</td>
<td>.0118</td>
<td>-.0638***</td>
</tr>
<tr>
<td></td>
<td>(.0194)</td>
<td>(.0174)</td>
</tr>
<tr>
<td>Constant</td>
<td>-1.43</td>
<td>-</td>
</tr>
<tr>
<td></td>
<td>(.113)</td>
<td></td>
</tr>
</tbody>
</table>

Observations  232  232  
(Pseudo) $R^2$ .009 .180

Table 1: Ordered Probit (1) and ordinary least squares (2) regression analyses predicting risk-attitude measures (Lottery Measure [LM] and Questionnaire Measure [GRF]) as a function of the Big Five personality factors (all subjects; $N = 232$). Table entries are the predictors’ un-standardised B-coefficients from multiple regression analyses; numbers in brackets indicate standard errors. *** $p < .01$, ** $p < .05$, * $p < .10$.

As a second step of our analysis, we compared the relationship between the different measures of individual risk attitudes and observed trusting behaviour. As pointed out earlier, the question about the correlation between risk taking and trusting behaviour is contentious. However, as we show below, the GRF provides indeed a powerful predictor not only for current but also for future (re-test) trusting behaviour.

16 The six decision domains considered by Nicholson et al. are: recreation, health, career, finance, safety and social.
thereby giving further support to a position arguing for a connection between risk and trust.

To begin with, we focus solely on the results of the test study, though. For this, the average transfer in the trust game was 4.664 Taler ($SD$ 3.095) with single transfers ranging from the full endowment of the truster (10 Taler) to none of it (0 Taler). Furthermore, only the questionnaire measure shows a significant correlation with investment behaviour in the trust game ($\rho=.151, p=.107$ for LM, and $\rho=.231, p=.013$ for GRF).

In order to further assess the predictive power of the two risk-attitude measures, we also ran several multiple ordinary least squares regression analyses trying to predict trusting behaviour by the lottery measure and the questionnaire measure (see Table 2).\textsuperscript{17} Controlling for gender, age and senders’ first-order beliefs regarding responders’ trustworthiness (in % of transferred amounts), and step-wise entering the lottery-and the questionnaire measure, respectively, actual trusting behaviour in fact turns out to be associated with both measures. However, the contribution of the lottery-choice measure is significant only at a 10% level and its inclusion into the regression adds only little to the explanatory power of the model. In particular, taking into account the results for the combined model (3) in which both LM and GRF are added, this strongly suggests that the questionnaire measure has considerably more predictive power regarding behaviour in the primary task (the trust game) than the lottery measure. And, in fact, the results of the re-test study presented below further support this interpretation of the results.

A possible explanation for the fact that the data from the lottery-choice task do not correlate with the general risk factor, personality or only weakly with actual risk-taking behaviour, of course, may be found in the fact that the lottery was also incentivised and was run after the trust-game. Thus, subjects may have used the lottery to hedge their behaviour in the trust game in a way that confounds the results of the lottery choice task (instead of treating the lottery-choice as an isolated item). Yet, we find no evidence for such behaviour in the data: As reported above, lottery choices are uncorrelated with investment behaviour. Moreover, there is also no significant correlation between the subjects’ beliefs about their expected outcome from the trust game and their lottery-choice behaviour. And, as we will report in more detail further below, there is also no significant correlation between the change of behaviour in the trust game and that in the lottery choice task over time. As also Houser et al.

\textsuperscript{17}Due to missing values (gender and age) we had to omit one subject.
Table 2: Ordinary least squares regression analyses predicting trusting behavior (Trust=amount transferred in trust game) as a function of LM, GRF, senders’ first order beliefs about back transfers (test, T1), controlling for gender and age; senders: N = 115. Table entries are the predictors’ un-standardised B-coefficients from multiple regression analyses; numbers in parentheses indicate robust standard errors. *** p < .01, ** p < .05, * p < .10.

(2010) find no order effects in a study on the determinants of trust, which combined the Berg et al. trust game with the Holt and Laury lottery-choice task in varying order, we are confident that order effects did not affect the outcomes in the present study.

### 3.2 The Re-Test

In order to investigate the test-re-test stability of both measures within a time period of one year, the re-test study was conducted under almost identical conditions (cf. Section 2.2); the results are reported below. In order to ensure that participants could not remember their prior responses, we asked them at the very end of the re-test study whether they had taken part in a similar study before. One participant could remember taking part in the first study and was therefore excluded from the below analyses leaving us with 22 trusters and 21 responders. Moreover, in view of the results it is noteworthy that the 44 participants who volunteered for phase 3 of the study did not, on either risk measure administered at phase 2, differ from those who did not (mean scores of 4.33 and 4.47 for risky lottery decisions, and .01 and -.02 for factor scores on the risk questionnaire, for phase 3 drop-outs and volunteers.

<table>
<thead>
<tr>
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<th>(2)</th>
<th>(3)</th>
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<td>0.242*</td>
<td>0.231*</td>
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</tr>
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<td></td>
<td>(0.124)</td>
<td>(0.125)</td>
<td></td>
</tr>
<tr>
<td>T1_GRF</td>
<td>0.669**</td>
<td>0.653**</td>
<td></td>
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<td></td>
<td>(0.260)</td>
<td>(0.266)</td>
<td></td>
</tr>
<tr>
<td>T1_FO-Belief</td>
<td>8.290***</td>
<td>8.114***</td>
<td>8.031***</td>
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<td></td>
<td>(1.309)</td>
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<tr>
<td>Constant</td>
<td>2.504*</td>
<td>3.632***</td>
<td>2.457*</td>
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<tr>
<td></td>
<td>(1.482)</td>
<td>(1.378)</td>
<td>(1.472)</td>
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<tr>
<td>Observations</td>
<td>115</td>
<td>115</td>
<td>115</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.316</td>
<td>0.336</td>
<td>0.354</td>
</tr>
</tbody>
</table>
respectively; Mann-Whitney Test, all \( p > .10 \) so that attrition should not have biased the reported results in favor of either measure of individual risk attitudes.\(^{18}\)

**Experimental Results - Re-Test Only**

All in all, the experimental results of the re-test study are very similar to those of the first study. The mean transfer was 4.409 Taler (\( SD \) 3.647) with transfers again ranging from the full endowment (10 Taler) to nothing (0 Taler). Moreover, as before, only the general risk factor is strongly correlated with trusting behaviour in the trust game (\( \rho = .494, p = .019 \)) while behaviour in the lottery-choice task is not (\( \rho = .238, p = .286 \)).

As above, we conducted several ordinary least squares regression analyses in order to examine the relation among the two measures of individual risk attitudes and observed trusting behaviour (see Table 3). This time, again holding gender, age and senders’ first-order beliefs constant, we only find the questionnaire measure to have some predictive power regarding actual trusting behavior in the re-test. Consistent with this, we find that controlling only for the lottery measure of risk attitudes, model (1), results in a predictive power of the model \( R^2 = .393 \) which is considerably lower than the predictive power of a model with only the general risk factor, model (2), \( R^2 = .597 \). Moreover, the complete model (3) again confirms this finding: only the general risk factor has significant predictive power regarding subjects’ trusting behaviour. Thus, the data again strongly suggest that the type of questionnaire measure advocated by Dohmen et al. (2011) is more suitable to assess the correlation of risk attitudes and in behaviour a the primary behavioural task in lab experiments, in this case the trust game (i.e. transfers in the trust game), than the lottery choice task.

**Test-Re-Test Comparison**

The main aspect of running the re-test study, however, was to scrutinise the over-time stability of the different measures elicited in the course of our experiment, i.e. the risk-attitude measures and the amount transferred in the trust game.

Regarding the stability of individual risk attitudes over time, we first analysed the correlations between the test and the re-test study responses. The general risk factor (as well as all single items of the underlying questionnaire) shows a very high over-time stability (\( \rho = .780, p < .001 \)), while the lottery-choice measure shows only a trend towards stability (\( \rho = .258, p = .095 \); the corresponding Pearson correlation

\(^{18}\)We also investigated demographic variables and trust decisions, but no differences between the groups were close to conventional levels of statistical significance; \( p > .10 \) for all items.
Table 3: Ordinary least squares regression analyses predicting trusting behavior in the re-test (T2_Trust=amount transferred in trust game in T2) as a function of LM, GRF, senders’ first order beliefs, controlling for gender and age; senders: N = 21. Table entries are the predictors’ un-standardised B-coefficients from multiple regression analyses; numbers in brackets indicate robust standard errors. *** p < .01, ** p < .05, * p < .10.

<table>
<thead>
<tr>
<th>VARIABLES</th>
<th>(1) T2_LM</th>
<th>(2) T2_GRF</th>
<th>(3) T2_FO-Belief</th>
<th>(4) Constant</th>
</tr>
</thead>
<tbody>
<tr>
<td>T2_LM</td>
<td>0.324</td>
<td>-0.449</td>
<td>0.324</td>
<td>-0.0566</td>
</tr>
<tr>
<td></td>
<td>(0.289)</td>
<td>(0.337)</td>
<td>(0.477)</td>
<td>(3.231)</td>
</tr>
<tr>
<td>T2_GRF</td>
<td></td>
<td>1.718***</td>
<td>2.422***</td>
<td>5.485***</td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.477)</td>
<td>(0.617)</td>
<td>(2.469)</td>
</tr>
<tr>
<td>T2_FO-Belief</td>
<td>10.89***</td>
<td>9.151***</td>
<td>8.396***</td>
<td>2.899</td>
</tr>
<tr>
<td></td>
<td>(3.130)</td>
<td>(2.602)</td>
<td>(2.533)</td>
<td>(2.469)</td>
</tr>
<tr>
<td>Constant</td>
<td>-0.0566</td>
<td>2.899</td>
<td>5.485***</td>
<td>5.485***</td>
</tr>
<tr>
<td></td>
<td>(3.231)</td>
<td>(2.469)</td>
<td>(1.784)</td>
<td>(1.784)</td>
</tr>
<tr>
<td>Observations</td>
<td>21</td>
<td>21</td>
<td>21</td>
<td>21</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.393</td>
<td>0.597</td>
<td>0.638</td>
<td></td>
</tr>
</tbody>
</table>

coefficient is not significant: $r = .205, p = .188$; see Table 4 for details and Figure 2 for illustration. Being surprised by the weak stability of choices in the lottery-choice task, we also examined the test-re-test stability of the lottery-choice measure by categorising participants into risk-averse (cross-over point after more than four safe decisions), risk-neutral (cross-over point after exactly four safe decisions), and risk-taking (cross-over point before more than four safe decisions). This could be done for 33 participants (those who had cross-over points in both studies). Yet, also the category shows no stability over time ($\chi^2 (4) = 5.38, p > .10$). Thus, the test-re-test comparison casts some further doubt on the reliability of the lottery-choice task as a measure of individual risk-attitudes in connection with other laboratory experiments.

For comparison purposes and to examine what type of test-re-test correlation that could reasonably be expected for a behavioural decision-making measure, we also computed the test-re-test correlation of the amount transferred in the trust game. This correlation was almost as high as that of the general risk factor ($\rho = .692, p < .001$) suggesting that trusting behaviour indeed has a component which is a stable individual characteristic as hypothesised by Glaeser et al. (2000, p. 827).
<table>
<thead>
<tr>
<th>Risk attitudes</th>
<th>Spearman’s $\rho$</th>
</tr>
</thead>
<tbody>
<tr>
<td>LM</td>
<td>$0.258^*$</td>
</tr>
<tr>
<td>GRF</td>
<td>$0.780^{***}$</td>
</tr>
<tr>
<td>General</td>
<td>$0.768^{***}$</td>
</tr>
<tr>
<td>Driving</td>
<td>$0.573^{***}$</td>
</tr>
<tr>
<td>Financial matters</td>
<td>$0.550^{***}$</td>
</tr>
<tr>
<td>Sport and leisure</td>
<td>$0.670^{***}$</td>
</tr>
<tr>
<td>Career</td>
<td>$0.623^{***}$</td>
</tr>
<tr>
<td>Health behaviour</td>
<td>$0.632^{***}$</td>
</tr>
<tr>
<td>Trusting strangers</td>
<td>$0.641^{***}$</td>
</tr>
</tbody>
</table>

Table 4: Test-re-test stability of risk measures (all subjects; $N = 43$). $^{***} p < .01$, $^{**} p < .05$, $^* p < .10$.

<table>
<thead>
<tr>
<th>VARIABLES</th>
<th>(1) T2_Trust</th>
<th>(2) T2_Trust</th>
<th>(3) T2_Trust</th>
</tr>
</thead>
<tbody>
<tr>
<td>T1_LM</td>
<td>0.0856</td>
<td>-0.00283</td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.414)</td>
<td>(0.351)</td>
<td></td>
</tr>
<tr>
<td>T1_GRF</td>
<td></td>
<td>1.374**</td>
<td>1.375**</td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.476)</td>
<td>(0.481)</td>
</tr>
<tr>
<td>T2_FO-Belief</td>
<td>10.59**</td>
<td>12.26***</td>
<td>12.27***</td>
</tr>
<tr>
<td></td>
<td>(3.936)</td>
<td>(2.169)</td>
<td>(2.983)</td>
</tr>
<tr>
<td>Constant</td>
<td>1.216</td>
<td>1.110</td>
<td>1.114</td>
</tr>
<tr>
<td></td>
<td>(3.161)</td>
<td>(2.890)</td>
<td>(2.961)</td>
</tr>
<tr>
<td>Observations</td>
<td>21</td>
<td>21</td>
<td>21</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.352</td>
<td>0.557</td>
<td>0.557</td>
</tr>
</tbody>
</table>

Table 5: Ordinary least squares regression analyses predicting trusting behavior in the re-test (T2) as a function of LM, GRF from the test study (T1), and senders’ first order beliefs (from the re-test study; T2) while controlling for gender and age ($N = 21$). Table entries are the predictors’ un-standardised B-coefficients from multiple regression analyses; numbers in brackets indicate robust standard errors. $^{***} p < .01$, $^{**} p < .05$, $^* p < .10$.

As a final step of our analysis regarding the reliability and over time stability of the two different measures of individual risk attitudes, we investigated the extent to which risk attitudes as measured in the test study can predict trusting behaviour in the re-test study (see Table 5); note that trusting behaviour in the re-test is not confounded
as it is not preceded by any other decision or questionnaire item. Consistent with our previous procedures, we used ordinary least squares regression analyses to try and predict investment decisions in the re-test by risk attitudes measures as gathered in the test study conducted one year earlier; again we controlled for gender, age and sender’s first-order beliefs in the re-test study. As it turns out, the questionnaire measure indeed has significant predictive power regarding actual trusting behavior in the re-test. The lottery measure, by contrast, shows no such effect but gives only insignificant coefficients.19

Figure 2: Decisions in the lottery-choice task (left panel) and risk questionnaire factor scores (right panel) on Times 1 (horizontal axis) and Time 2 (vertical axis). OLS regression lines (betas = .2252 and .6818, for left and right panel, respectively) fitted to the data.

Testing for Confounding Order Effects.
Finally, we used the results of the re-test study for a further test of the possibility that the results of the lottery-choice task are confounded due to order effects, i.e. a conditioning of lottery choices on investment behaviour in the trust game that might have blurred the underlying correlation between risk and trust. In particular, we tested whether the direction of over-time changes in investment behaviour (trust

19Single correlations of risk attitude measures from the test study and trust transfers in the re-test yield no significant correlation coefficients.
game) are correlated with direction of changes in lottery-choice behaviour. If subjects
had conditioned their behaviour in the lottery choice task on their earlier investment
behaviour, it seemed reasonable to expect the direction of the deviations to be corre-
related. However, we do not find any evidence for such a correlation: 11 trusters change
their investments over time (3 invest more, 8 less). Of those who invest more in the
re-test-study, one makes a safer choice in the lottery task and two make a more risky
one, while the 8 subjects who invest less in the re-test-study split equally into safer
and more risky choices (4 each). Accordingly, we find that the direction of change
in the trust game is not correlated with the direction of change in the lottery-choice
task (Fisher’s exact test $p = .576$). Thus, also the combination of test and re-test
data gives no indication of order effects which might have confounded the results of
the lottery-choice task.

4 Concluding Remarks

In this paper, we have compared the empirical power of two different measures of indi-
vidual risk attitudes as instruments to analyse behaviour in economic lab experiments
(here the Berg et al., 1995, trust-game): the Holt and Laury (2002) lottery choice task
(with small stakes) and a questionnaire similar to the one advocated by Dohmen et al.
(2011). The data presented above show that the general risk factor derived from the
questionnaire about individual risk attitudes has good construct validity (being corre-
lated with an external predictor of risk-taking behaviour, namely personality, almost
exactly as expected), reasonable predictive power (regarding transfer behaviour in the
trust game), and a very good test-re-test stability over time (one year). By contrast,
the Holt and Laury (2002) lottery-choice measure shows no construct validity (again
when related to personality), almost no predictive power, and most importantly no
robust test-re-test stability.21 Furthermore, we also find no evidence supporting the

20The same remains to hold if also trustees are drawn into the picture: 10 of them change their
back-transfer behaviour over time. Combining those who invest more (less) with those who back-
transfer more (less), we again find that the direction of change in one task is not correlated with
that in the other (Fisher’s exact test $p = .410$).

21Note that we are not the first to study over-time stability of individual risk attitudes. In a
different context, using a different subject pool (a representative sample of the Danish population)
and a different set-up (for different scalings of the lottery with no further primary experimental
task), for example, Harrison, Johnson, McInnes and Rutström (2005b) find that CRRA coefficients
when inferred from choices in the Holt and Laury lottery-choice measure exhibit some stability over
time; see also Andersen, Harrison, Lau and Rutström (2008b). We are not aware of any other study
analysing and comparing the test-retest stability of (different) measures of individual risk attitudes
in connection with other primary tasks in economic lab experiments, though.
convergent validity between general risk factor and the Holt and Laury (2002) lottery-choice task, i.e. both are uncorrelated in our study. While we are hesitant to draw any bold conclusion as to source of these findings, we nonetheless interpret our results as suggesting that behaviour in the lottery-choice task is a comparably unreliable measure of individual differences in risk-attitudes— an observation which should at least be kept in mind whenever risk-attitudes are measured in connection with other (primary) behavioural tasks in the lab.

Moreover, it is noteworthy that the lack of reliability of the lottery-choice measure can neither be clearly attributed to the method of measurement (i.e. relying on a single behavioural item), as our other one-shot behavioural measure – transferred amount as truster – shows a very high test-re-test stability. Nor can it be attributed to the construct being measured (i.e. risk attitudes), as the multi-item questionnaire measure shows a very high test-re-test stability. Thus, the comparably high degree of noise appears to be specific to the behavioural measurement of individual risk attitudes by means of lottery choice tasks (with small stakes) such as the one proposed by Holt and Laury (2002). From an applied point of view, our results therefore suggest using the type of questionnaire measure advocated by Dohmen et al. (2011) rather than — or at least complementary to — common lottery-choice measures when studying the connection between individual risk-attitudes and behaviour in laboratory experiments; this holds in particular if stable qualitative individual differences in attitudes towards risk are important rather than information about the exact curvature of the utility functions.

Apart from the aspects about the adequate measurement of individual risk attitudes mentioned above, our results also contribute to the current discussion on the relation between risk and trust. Challenging the claim that there is an intrinsic relation between these constructs (suggested e.g. by Snijders and Keren, 1998, or Ben-Ner and Putterman, 2001), Eckel and Wilson (2004) reported that their risk instruments (including the Holt and Laury measure) could not predict the decision to trust in a binary trust game (for similar results, see Bohnet and Zeckhauser, 2004; Houser et al., 2010). By contrast, our results support the intuitive contention that trusting behaviour, as revealed in the Berg et al. trust / investment game, does entail an element of risk. In how far these results translate to more general instances of trust,

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22 There might be fundamental problems connected to lottery-choice measures giving rise to these findings or simply differences in the cognitive ability of the subjects as suggested, for example, by Dave et. al. (2010) or Anderson et al. (2013).

23 The more general claim going beyond the Holt and Laury task, of course, calls for further research to be substantiated.
of course, remains an open question. The results of the present study, however, strongly suggest that future research into the nature of trust and its connection to risk should include a questionnaire measure of individual risk-attitudes of the type advocated by Dohmen et al. (2011) instead of relying on lottery-choice tasks alone.

Finally, we want to reemphasise that the high test-re-test stability of the observed investment behaviour in the trust game strongly suggests that trusting behaviour indeed has a component which is a stable individual characteristic as hypothesised by Glaeser et al. (2000, p. 827). Thus, although the exact determinants of trust remain an open question, it seems that trust is at least a comparably stable phenomenon — a fact that, despite all remaining difficulties, should facilitate future enquiries into the nature of trust.

References


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24See Fehr (2009) for an interesting recent discussion about the determinants of trust.

25As mentioned earlier, the respective hypothesis is also supported through a study by Cesarini et al. (2008) who identify a correlation between trusting behaviour and the subjects’ genetic code.


