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Compensated Discrete Choice with Particular Reference to Labor Supply

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Compensated Discrete Choice with Particular Reference to Labor Supply

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Compensated Discrete Choice with Particular Reference to Labor Supply

Abstract

Dagsvik and Karlström (2005) have demonstrated how one can compute Compensating Variation and Compensated Choice Probabilities by means of analytic formulas in the context of discrete choice models. In this paper we offer a new and simplified derivation of the compensated probabilities. Subsequently, we discuss the application of this methodology to compute compensated labor supply responses (elasticities) in a particular discrete choice labor supply model. Whereas the Slutsky equation holds in the case of the standard microeconomic model with deterministic preferences, this is not so in the case of random utility models. When the non-labor income elasticity is negative the Slutsky equation implies that the compensated wage elasticity is higher than the uncompensated one. In contrast, in our random utility model we show empirically that in a majority of cases the uncompensated wage elasticity is in fact the highest one. We also show that when only the deterministic part of the utility function is employed to yield optimal hours and related elasticities, these elasticities are numerically much higher and decline more sharply across deciles than the random utility ones.

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Keywords: female labor supply, compensated discrete choice.

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1. Introduction

In recent years labor supply analysis based on the theory of discrete choice and random utility formulations has become increasingly popular. A major reason is that discrete choice labor supply models are much more practical than the conventional continuous approach based on marginal calculus: see the surveys by Creedy and Kalb (2005) and Dagsvik et al. (2013). For example, with the discrete choice approach, it is easy to deal with non-linear and non-convex economic budget constraints, and to apply rather general functional forms of the utility representations. In the literature basically two versions of discrete models of labor supply have been proposed. Van Soest (1995) and van Soest et al. (2002) propose to analyzing labor supply as a standard discrete choice problem. Alternatively, Dagsvik (1994), Aaberge et al. (1995), Aaberge et al. (1999), with further work by Dagsvik and Strøm (2006), Dagsvik and Jia (2006) and Dagsvik and Jia (2013), propose to analyzing labor supply as a job choice problem, where the set of feasible jobs is individual-specific and latent.

The history of random utility models dates back at least to Thurstone's 1927 analysis. It was more than 30 years before these types of models were applied in economics: see Quandt (1956), Luce (1959), and McFadden (1973, 1978, 1981, 1984) for wide-ranging discussions of motivation, exposition, and application. An early analysis of the potential of the random utility approach within economics was provided by Quandt (1956). He confronted the established wisdom of the time by arguing that it might be better to define preferences in a probabilistic sense, thereby making the model more realistic by weakening the postulate of strict rationality.

In microeconomics, the theory of compensated (Hicksian) demand and supply plays an important role. Until recently there has been very little focus on a corresponding theory for random utility models. While it is fairly straightforward to compute uncompensated responses and elasticities in discrete choice models, computation of the corresponding compensated effects is not so easy. Dagsvik and Karlström (2005) have demonstrated how one can calculate compensating variation (CV) and compensated choice probabilities (Hicks choice probabilities) in random utility models in the general case where the deterministic part of the utility function may be non-linear in income. Kornstad and Thoresen (2006) and Dagsvik et al. (2009) have applied the methodology of Dagsvik and Karlström (2005) to compute welfare measures such as CV in the context of selected tax reforms.

The contribution of this paper is twofold. First, we show that the derivations of compensated probabilities and CV given by Dagsvik and Karlström (2005) can be substantially simplified in the case of random utility models with utilities that are non-linear in income. Second, we show how one can apply the methodology of Dagsvik and Karlström (2005) to compute Hicks choice probabilities and corresponding compensated elasticities. To this end we base our analysis on the model developed by Dagsvik and Strøm (2006). Specifically, we show how one can compute Hicks joint choice probabilities and corresponding elasticities of being in particular states before and after tax reforms or changes in wages. The states are non-working, working in specific sectors and with different working loads.

We report uncompensated as well as compensated wage elasticities, and also non-labor income elasticities. These calculations show that the compensated wage elasticities tend to be higher at the extensive margin than at the intensive margin, while the opposite is true for uncompensated wage elasticities. The elasticities vary considerably with wage levels (the higher the wage, the lower the elasticity) and with household characteristics. Moreover, the calculation of these three types of elasticities demonstrates the extent to which the Slutsky equation is violated for this type of model.

Finally, we calculate compensated elasticities based on the sample values that were used in estimating the model. As in the case above, we find that the mean uncompensated elasticities of expected hours with respect to wages exceed the compensated elasticities, which is inconsistent with the prediction from the Slutsky equation.

The paper is organized as follows. In the next section we provide a new and simplified version of some of the key results obtained by Dagsvik and Karlström (2005) regarding compensating variation and compensated choice probabilities. We then discuss special cases, such as binary and ternary choice settings. This will demonstrate the essential features of the methodology proposed by Dagsvik and Karlström (2005), especially in relation to the labor supply. In Section 3 we provide a brief description of the empirical labor supply model that is the main focus of our application and in Section 4 we give numerical results.

2. Expenditure and compensated choice in random utility models

In this section we review a fundamental result obtained by Dagsvik and Karlström (2005) and we provide a simplified proofs.

Consider a setting with a finite number of alternatives. Assume that the utility of alternative j has the structure $U_j(y) = \log v_j(y) + \varepsilon_j$, where $v_j(y)$ is a positive, deterministic and monotonically increasing function of y and may also depend on prices, non-pecuniary and alternative j - specific attributes, whereas ε_j is a stochastic term that is supposed to account for the effect on preferences from variables that are not observed by the researcher.⁴ To this end let $U_j^0 = \log v_j^0 + \varepsilon_j = \log v_j^0(y^0) + \varepsilon_j$ denote the ex-ante utility of alternative j , where y^0 denotes the initial income and $v_j^0(\cdot)$ is the deterministic term associated with the utility of alternative j ex-ante, and let $U_j(y) = \log v_j(y) + \varepsilon_j$ be the corresponding utility of alternative j ex-post. The respective systematic terms $v_j^0(\cdot)$ and $v_j(\cdot)$ may differ due to exogenous changes in taxes or alternative specific attributes, suppressed in the notation here. Here it is assumed that the stochastic terms of the utility function are not affected by the reform. Let J^0 and J denote the ex-ante and ex-post choice given that the ex-ante and ex-post

⁴ With no loss of generality we apply the logarithm transformation for the sake of simplifying the expressions below.

utility levels of the chosen alternatives are equal. Let $P^H(j, k)$ denote the joint compensated (Hicksian) probability of choosing alternative j ex-ante and alternative k ex-post under the condition that the respective utility levels of the chosen alternatives before and after the reform are the same. In other words,

$$(2.1) \quad P^H(j, k) = P(U_j^0 = \max_r U_r^0, U_k(Y) = \max_r U_r(Y), \max_r U_r^0 = \max_r U_r(Y)),$$

where Y is the expenditure function that yields the income required to maintain the utility level equal to the original utility level. Although suppressed in the notation, it is implicit that the expenditure function depends on the ex ante utility level as well as ex ante and ex post prices and other attributes. Furthermore, let

$$(2.2) \quad \varphi(j, k, y) = P(U_j^0 = \max_r U_r^0, U_k(Y) = \max_r U_r(Y), \max_r U_r^0 = \max_r U_r(Y), Y \leq y).$$

The interpretation of the expression in (2.2) is as the joint probability of choosing alternative j ex-ante, alternative k ex-post, and $\{Y \leq y\}$, when the ex-post maximum utility is equal to the ex-ante maximum utility.

Note that the income Y is stochastic due to the utility function containing a stochastic term. Let y_j be determined by $v_j^0(y^0) = v_j(y_j)$. That is, y_j is the ex-post income that ensures that the ex-ante utility and ex-post utility of alternative j are equal. Also let $\psi_r(y) = \max(v_r^0(y^0), v_r(y))$. For simplicity we shall write $v_j^0(y^0) = v_j^0$. The following result was obtained by Dagsvik and Karlström (2005, Corollary 3).

Theorem 1

Assume an additive random utility model with independent error terms that have c.d.f. $\exp(-\exp(-x))$.

Then

$$(2.3) \quad \varphi(j, k, dy) = \frac{v_j^0 v_k(dy)}{\left\{ \sum_{r=1}^m \psi_r(y) \right\}^2}$$

provided that $k \neq j$ and $y_j \geq y \geq y_k$.⁵

The proof of Theorem 1 is given in Appendix A. As mentioned above, the proof given in Appendix A is a much simpler proof than one given by Dagsvik and Karlström (2005).

For the sake of demonstrating the essential idea, we shall provide a brief outline of the argument of the proof: Consider the case with $m = 3$. We have

$$\{J^0 = 1, J = 2\} \Leftrightarrow \{\max(U_2^0, U_3^0) \leq U_1^0, \max(U_1(Y), U_3(Y)) \leq U_2(Y) = U_1^0\}.$$

For j to be the most preferred alternative ex-ante and k the most preferred alternative ex-post, one must have $U_2(Y) = U_1^0 > U_2^0$. Furthermore, since alternative 2 is the most preferred one ex-post, $U_2(Y) > U_1(Y)$, which

⁵ As usual, the notation dy means $(y, y + dy)$.

implies that $U_1(Y) < U_1^0$. Hence, the event $\{U_2(y) = U_1^0\}$ implies that $\{U_1(y) \leq U_1^0, U_2^0 \leq U_1^0\}$. Accordingly, the formula above yields

$$\{J^0 = 1, J = 2, Y = y\} \Leftrightarrow \{\max(U_3^0, U_3(y)) \leq U_1^0 = U_2(y)\}.$$

The last equation represents a key relation for deriving corresponding probabilities. As shown in Appendix A one can now readily calculate

$$\begin{aligned} &= P(\max(U_3^0, U_3(y)) \leq U_1^0, U_2(y) \leq U_1^0 \leq U_2(y + \Delta y) | U_1^0 = x) \\ &= P(\max(U_3^0, U_3(y)) \leq x, U_2(y) \leq x \leq U_2(y + \Delta y)) \end{aligned}$$

where Δy is a small positive value. Note that the expression above is completely general and does not depend on the assumption of independent extreme value distributed random terms, so it is not hard to see how one can extend the result of Theorem 1 to the case with correlated error terms. In the case with independent error terms one has that

$$P(\max(U_3^0, U_3(y)) \leq x, U_2(y) \leq x \leq U_2(y + \Delta y)) = P(\max(U_3^0, U_3(y)) \leq x)P(U_2(y) \leq x \leq U_2(y + \Delta y))$$

which simplifies the subsequent calculations further. By using the relation

$$P(J^0 = 1, J = 2, Y \in [y, y + dy]) = EP(\max(U_2^0, U_3^0) \leq U_1^0, \max(U_1(Y), U_3(Y)) \leq U_2(Y), Y \in [y, y + \Delta y) | U_1^0)$$

one can obtain the desired choice probabilities.

The results in the following corollaries are also given in Dagsvik and Karlström (2005), but the proofs given here are considerably simpler.

Corollary 1

Under the assumptions of Theorem 1 the Hicksian (compensated) choice probabilities are given by

$$(2.4) \quad P^H(j, k) = v_j^0 \int_{y_k}^{y_j} \frac{v_k(dy)}{\left\{ \sum_{r=1}^m \psi_r(y) \right\}^2}$$

when $k \neq j$ and $y_j > y_k$. In the case where j and k are distinct and $y_j < y_k$ then $P^H(j, k) = 0$. Furthermore, when $j = k$, then

$$(2.5) \quad P^H(j, j) = \frac{v_j^0}{\sum_{r=1}^m \psi_r(y_j)}.$$

The proof of Corollary 1 is given in Appendix A.

Corollary 2

Under the assumptions of Theorem 1 the probability density, $\gamma(x)$ of Y , is given by

$$\gamma(y) = \sum_{j,k} 1\{y_j \geq y \geq y_k\} \partial \varphi(j, k, y) / \partial y = \frac{\sum_j 1\{y_j \geq y\} v_j^0 \sum_{r=1}^m v'_r(y)}{\left\{ \sum_{r=1}^m \psi_r(y) \right\}^2}$$

provided the deterministic parts of the utility function are continuously differentiable.

The proof of Corollary 2 is given in Appendix A.

The corresponding *Compensating Variation* (CV) follows from the expenditure function by the relation $CV = y^0 - Y$. From the results above one can obtain the corresponding c.d.f. of the expenditure function.

Corollary 3

Under the assumptions of Theorem 1 with $y < y_j$, we have

$$P(J^0 = j, Y > y) = \frac{v_j^0}{\sum_{r=1}^m \psi_r(y)}.$$

The proof of Corollary 3 is given in Appendix A.

Dagsvik and Karlström (2005) prove the result of Corollary 3 (their Theorem 3) by a direct argument. It may now be instructive to look at a couple of special cases, namely the binary and ternary cases.

2.1. The binary case

In order to provide intuition into the formalism developed above, we shall now consider the case with two alternatives (alternatives one and two) in more detail. From (2.2) it follows by straightforward integration that when $y_1 > y_2$

$$(2.6) \quad P^H(1, 2) = v_1^0 \int_{y_2}^{y_1} \frac{v_2(dy)}{\{v_1^0 + v_2(y)\}^2} = \frac{v_1^0}{v_1^0 + v_2^0} - \frac{v_1^0}{v_1^0 + v_2(y_1)}.$$

According to (2.4), $P^H(2, 2)$ is given by

$$(2.7) \quad P^H(2, 2) = \frac{v_2^0}{\sum_{k=1}^2 \psi_k(y_2)} = \frac{v_2^0}{v_1^0 + v_2^0}.$$

Consequently, it follows that the ex post compensated probability of choosing alternative 2 is given by

$$(2.8) \quad P_2^H = P^H(1,2) + P^H(2,2) = \frac{v_2(y_1)}{v_1^0 + v_2(y_1)}$$

in this case. The corresponding ex-ante probability of choosing alternative 2 is given by

$$(2.9) \quad P_2 = \frac{v_2^0}{v_1^0 + v_2^0}.$$

Thus it follows that the compensated effect, when measured by the relative change $(P_2^H - P_2)/P_2$, is given by

$$(2.10) \quad \begin{aligned} \frac{P_2^H - P_2}{P_2} &= \left(\frac{v_2(y_1)}{v_1^0 + v_2(y_1)} - \frac{v_2^0}{v_1^0 + v_2^0} \right) / P_2 \\ &= \frac{v_2(y_1) - v_2^0}{v_1^0 + v_2(y_1)} = \frac{v_2(y_1) - v_2(y_2)}{v_1^0 + v_2(y_1)} > 0 \end{aligned}$$

when $y_1 > y_2$.

In the case where $y_1 \leq y_2$, $P^H(1,2) = 0$ and $P_2^H = P^H(2,2)$, so that we now obtain

$$(2.11) \quad P_2^H = P^H(2,2) = \frac{v_2^0}{v_2^0 + v_1(y_2)}.$$

Hence, in the case with $y_1 \leq y_2$,

$$(2.12) \quad \frac{P_2^H - P_2}{P_2} = \frac{v_1^0 - v_1(y_2)}{v_2^0 + v_1(y_2)} = \frac{v_1(y_1) - v_1(y_2)}{v_2^0 + v_1(y_2)} \leq 0.$$

Consider next the special case with only a reduction in the price (or cost) of alternative 2 while other attributes remain fixed. This implies that $v_1(y) = v_1^0(y)$ in this case, so that $y_1 = y^0$. Since $v_2(y) > v_2^0(y)$ for any y , and y_2 is determined by $v_2(y_2) = v_2^0(y^0)$, it must be the case that $y_2 < y^0 = y_1$. According to (2.12), this means that the compensated price elasticity $(P_2^H - P_2)/P_2$ can never be negative in this case, which is intuitive.

2.2. The ternary case

Consider finally the case with three alternatives: that is, $m = 3$. Then (2.4) reduces to

$$(2.13) \quad P^H(j,k) = v_j^0 \int_{y_k}^{y_j} \frac{v_k(dy)}{\left\{ \sum_{r=1}^3 \psi_r(y) \right\}^2},$$

for $j, k = 1, 2, 3$, and distinct j and k . Suppose, for example, that $y_1 > y_3 > y_2$. It then follows that $P^H(1,2) > 0$, $P^H(1,3) > 0$, $P^H(3,2) > 0$, whereas $P^H(2,3) = P^H(3,1) = P^H(2,1) = 0$. From this we obtain

$$\begin{aligned}
(2.14) \quad P^H(3,2) &= v_3^0 \int_{y_3}^{y_1} \frac{v_2(dy)}{\{v_1^0 + v_2(y) + v_3^0\}^2} \\
&= \frac{v_3^0}{v_1^0 + v_2(y_3) + v_3^0} - \frac{v_3^0}{v_1^0 + v_2(y_1) + v_3^0} \\
&= \frac{v_3^0(v_2(y_1) - v_2(y_3))}{(v_1^0 + v_2(y_3) + v_3^0)(v_1^0 + v_2(y_1) + v_3^0)}.
\end{aligned}$$

However, one cannot in general not express the integral in (2.13) on closed form. The ex-post Hicksian probability of choosing alternative 2 equals

$$(2.15) \quad P_2^H = P^H(1,2) + P^H(3,2) + P^H(2,2),$$

and similarly for the other cases. Since $P^H(3,1) = P^H(2,1) = 0$, we get for alternative 1 that

$$(2.16) \quad P_1^H = P^H(1,1) = \frac{v_1^0}{v_1^0 + v_2(y_1) + v_3(y_1)}.$$

2.3. Examples of compensated effects in discrete labor supply models

Consider now the case of a model for labor force participation (binary case, $j=1,2$) and subsequently a model for labor force participation and choice of working in one of two sectors (ternary case, $j=1,2,3$).

Example 1: Labor force participation (the binary case)

Consider first the choice model of whether or not to work. Let w and I be the agent's wage and non-labor income. Let $f(hw, I)$ denote income after tax, where $h = 1$ if working and $h = 0$ otherwise. Let

$U_1(y) = u_1(f(0, y)) + \varepsilon_1$ be the agent's utility of not working and $U_2(y) = u_2(f(w, y)) + \varepsilon_2$, the utility of working, where $u_j(C)$ is a deterministic function and the random error term ε_j , $j = 1, 2$, are supposed to

account for unobserved heterogeneity in preferences across alternatives and agents. The agent will work if $U_2(y) > U_1(y)$. In this example it is assumed that the agent has no problem with finding a job, and consequently being employed will therefore be determined solely by the agent's preferences.

Consider now the effect of a change in economic incentives (such as a wage increase or a tax reform) that makes the working alternative more (or less) attractive. As above we assume in the following that the random part of the utility function is kept fixed under the reforms. In other words, the results are to be interpreted as conditional on the unobservable. Let w^0, f^0 be the ex-ante wage and tax system, and w, f , the corresponding wage and tax system ex-post. Moreover, let Y be the non-labor income that makes the ex-post indirect utility equal to the ex-ante indirect utility. Using the notation in section 2.1, let

$\log v_j^0(y^0) = \log v_j^0 = u_j(f^0((j-1)w^0, y^0))$, $U_j^0 = \log v_j^0(y^0) + \varepsilon_j$, $\log v_j(Y) = u_j(f((j-1)w, Y))$, and $U_j(Y) = \log v_j(Y) + \varepsilon_j$, for $j = 1, 2$. Thus U_j^0 is the ex-ante utility of being in labor market state j and $U_j(Y)$ the corresponding ex-post utility. The random parts, ε_j , $j = 1, 2$, are assumed i.i. extreme value distributed across alternatives. The deterministic utilities may also depend on other alternative specific attributes. Let y_j be defined by $v_j^0(y^0) = v_j(y_j)$, $j = 1, 2$. The compensated probability of “working” (“not working”) can now be analyzed as outlined above in Section 2.1. In particular, it follows that the compensated wage elasticity will never be non-positive.

Example 2: A two-sector labor supply model (the ternary case)

Consider next the following two-sector labor supply model. In this case there are three alternatives: “not working” (1), “working in the public sector” (2), and “working in the private sector” (3). Similarly to the previous example the only working option in both sectors is to work full time. As above, let U_1 be the utility of not working, U_2 the utility of working in the public sector and U_3 the utility of working in the private sector. Assume that $U_1(y) = u_1(f(0, y)) + \varepsilon_1$ and $U_j(y) = u_j(f(w_j, y)) + \varepsilon_j$ for $j = 2, 3$, where w_j is the wage of sector j . Let $\log v_1(y) = u_1(f(0, y))$ and $\log v_j(y) = u_j(f(w_j, y))$, for $j = 2, 3$. As with the analysis in the previous example, let $U_j^0 = \log v_j^0(y^0) + \varepsilon_j$ be the ex-ante utility representation and $U_j(Y) = \log v_j(Y) + \varepsilon_j$ the corresponding ex-post utility representation. The analysis now proceeds in a similar way to the ternary case discussed above.

3. Compensated choice in a labor supply model with latent choice sets of available jobs

In this section we consider compensated choice probabilities for the discrete labor supply model for married women with latent choice constraint developed by Dagsvik (1994) and Dagsvik and Strøm (2006), see also Dagsvik et al. (2013). The two-sector labor supply model is outlined in Appendix B. The wage income of the husband is assumed to be exogenously given. The household derives utility from household consumption, here set equal to household disposable income, leisure and non-pecuniary latent attributes of jobs.

Let $z = 1, 2, \dots$, be an indexation of the jobs and let $z = 0$ represent not working. The utility function is assumed to have the form

$$(3.1) \quad U(C, h, z) = \log v(C, h) + \varepsilon(z),$$

where (C, h) denotes disposable income and annual hours of work and $v(\cdot)$ is a suitable positive deterministic function. Let $B(h)$ be the set of jobs with hours of work h that are available to the agent. Let D be the set of feasible hours, which here is assumed to be finite. The sets $B(h)$, $h \in D$, are individual-specific and latent. The terms $\{\varepsilon(z)\}$ are positive sector- and job-specific random taste shifters. The taste shifter accounts for unobserved individual characteristics and unobserved job-specific attributes. These taste shifters, $\{\varepsilon(z)\}$, are assumed to be i.i.d. across jobs and agents, with c.d.f. $\exp(-e^{-x})$, for real x . The reason why the index z enters the utility function is that job-specific attributes beyond wage and hours of work may affect the utility of the agents. For given hours of work h and wage rate w , disposable household income is given by $C = f(hw, I)$, where $f(\cdot)$ is a function that transforms pre-tax incomes into after-tax incomes, w is the woman's wage and I denotes non-labor income. Let θ be a measure that represents the number of jobs available to the agent and $g(h)$ the fraction of jobs with hours of work that are available to the agent. Thus $\theta g(h)$ is the number of jobs in the latent set $B(h)$. We call $\theta g(h)$ the opportunity measure of the agent. From (3.1) it follows that the highest utility the agent can attain, given that hours of work are equal to h , is obtained by

$$(3.2) \quad V(h, y) = \max_{z \in B(h)} U(f(hw, y), h, z) = \log v(f(hw, y), h) + \max_{z \in B(h)} \varepsilon(z) = \log(v(f(hw, y), h) \theta g(h)) + \eta(h)$$

where the error terms $\{\eta(h)\}$ have the same distributional properties as $\{\varepsilon(z)\}$. The last equality in (3.2) follows because the distributional assumptions about the random error terms imply that

$$\eta(h) = \max_{z \in B(h)} \varepsilon(z) - \log(\theta g(h))$$

has the same distribution as $\varepsilon(1)$. This property follows readily from the extreme value distributional assumption, but to make the paper self-contained a proof of this property is given in Appendix A. Thus, formally, the utility maximization (with respect to hours of work) in the presence of this type of latent constraint can be achieved from the corresponding unconstrained case by modifying the structural part of the utility function, $v(f(hw, I), h)$, by multiplication of $\theta g(h)$.

Let $\varphi(h)$ be the probability (uncompensated) of choosing hours of work h (for a utility maximizing agent). From (3.2) it follows immediately from the theory of discrete choice that

$$(3.3) \quad \varphi(h) = P(V(h, I) = \max_{x \in D \cup \{0\}} V(x, I)) = \frac{v(f(hw, I), h) g(h) \theta}{v(f(0, I), 0) + \sum_{x > 0, x \in D} v(f(xw, I), x) g(x) \theta}$$

for $h > 0$, see Dagsvik and Strøm (2006) or Dagsvik et al. (2013). For $h = 0$, $\varphi(0)$ is obtained from (3.3) by replacing the numerator by $v(f(0, I), 0)$. Dagsvik and Jia (2013) discuss identification of this type of models and they show that the model is identified under the functional form assumptions described in Appendices A and B.

As in Section 2, consider a setting where a tax reform, a wage change, or some other change is introduced. For example, the framework above allows for changes in latent choice constraints through the

opportunity measure $\theta g(h)$. Let $(f^0, w^0, \theta^0 g^0(h))$ and $(f, w, \theta g(h))$ denote the ex-ante and ex-post tax system wages and opportunity measures respectively. The corresponding compensated effects can now be calculated in a similar way to that outlined in Section 2. Let $y(h)$ be the real number that solves

$$v(f^0(hw^0, I), h)\theta^0 g^0(h) = v(f(hw, y(h)), h)\theta g(h)$$

when h is positive and

$$v(f^0(0, I), 0) = v(f(0, y(0)), 0)$$

when $h = 0$. For short, let $V^0(h, y)$ and $V(h, y)$ denote the ex-ante and ex-post maximal utility given hours of work and non-labor income (h, y) . Define the joint compensated probability

$$(3.4) \quad P^H(h^0, h) = P(\max_{x \in D \cup \{0\}} V^0(x, I) = V^0(h^0, I), V(h, Y) = \max_{x \in D \cup \{0\}} V(x, Y), \max_{x \in D \cup \{0\}} V^0(x, I) = \max_{x \in D \cup \{0\}} V(x, Y)).$$

The probability defined in (3.4) is similar to the corresponding one defined in Section 2. It is the probability that the ex-ante labor supply is equal to h^0 and the ex-post supply is equal to h , when the ex-ante and ex-post utility levels are equal. It follows from Corollary 1 that

$$(3.5) \quad P^H(h^0, h) = 1\{y(h^0) > y(h)\} \int_{y(h)}^{y(h^0)} \frac{\theta^0 g^0(h^0)\theta g(h)v(f^0(h^0 w^0, I), h^0)v(f(hw, dy), h)}{K(y)^2},$$

for $h \neq h^0$, and $h > 0$, $\tilde{h} > 0$, where the indicator notation $1\{\cdot\}$ means that $1\{y > x\}$ if $y > x$ and zero otherwise and

$$K(y) = \max(v(f^0(0, I), 0), v(f(0, y), 0)) + \sum_{x \in D} \max[\theta^0 g^0(x)v(f^0(xw^0, I), x), \theta g(x)v(f(xw, y), x)].$$

Recall that the choice probabilities in (3.1) and (3.5) allow for changes in the opportunity measure, where $\theta^0 g^0(h)$ and $\theta g(h)$ denote the ex-ante and ex-post opportunity measures respectively. Furthermore, we get from Corollary 1 that

$$(3.6) \quad P^H(h, h) = \frac{\theta^0 g^0(h)v(f^0(hw^0, I), h)}{K(y(h))},$$

for $h > 0$. For $h^0 = 0$, $h > 0$, it follows that

$$(3.7) \quad P^H(0, h) = 1\{y(0) > y(h)\} \int_{y(h)}^{y(0)} \frac{\theta g(h)v(f^0(0, I), 0)v(f(hw, dy), h)}{K(y)^2},$$

and similarly for $h^0 > 0$, $h = 0$. For $h = h^0 = 0$, we have

$$(3.8) \quad P^H(0,0) = \frac{v^0(f^0(0,I),0))}{K(y(0))}.$$

Let $\varphi^0(h)$ denote the ex-ante choice probability, which is obtained from (3.3) by inserting the initial tax system, wage, and opportunity distribution. The marginal ex-post compensated choice probability, $\varphi^H(h)$ is given by

$$(3.9) \quad \varphi^H(h) = \sum_{x \in D} P^H(x, h).$$

Hence the compensated relative change in the probability of choosing h hours of work induced by the reform equals

$$(3.10) \quad \frac{\varphi^H(h) - \varphi^0(h)}{\varphi^0(h)}.$$

The corresponding compensated change in the mean labor supply is given by

$$(3.11) \quad \frac{\sum_{x \in D} x \varphi^H(x) - \sum_{x \in D} x \varphi^0(x)}{\sum_{x \in D} x \varphi^0(x)}.$$

4. Computation of selected compensated and uncompensated effects

Here we report numerical results based on the two-sector discrete labor supply model developed by Dagsvik and Strøm (2006). In Appendix B we give a summary of the model. Summary statistics, tax functions, and estimates are given in Appendix C.

4.1. Elasticities

We have applied the model to compute compensated wage elasticities and uncompensated wage and income elasticities for 48 different cases. We have chosen three wage levels (NOK 1994): one that is low (NOK 70 per hour), one that is high (NOK 200 per hour), and one that is super-high (NOK 300 per hour). Non-labor income (which include the income of the husband) are NOK 50,000 (low), NOK 100,000 (lower than average), and NOK 200,000 (above average). The selected household characteristics are: “No child” or “two children”, combined with woman’s age equal to 30 years or 40 years. The results below demonstrate the extent of heterogeneity in the labor supply responses. Further detailed results can be found in Appendix D. The levels of participation and expected hours of work prior to the wage increase are given in Appendix E. The tables in Appendix E show that participation and hours worked, given the wage decline with non-labor income, number of children, and age. Hours worked increase with the wage level.

The elasticities vary to a great extent. The highest uncompensated elasticity of conditional expected hours with respect to the wage is above 4 (public sector) and 2.5 (the private sector) times higher than the lowest elasticity (Table 1). For the uncompensated elasticity of unconditional hours the ratios are around 8.5 (public

sector) and 6 (private sector). The differences between the highest and lowest compensated elasticities are smaller (Table 2) than for the uncompensated. Table 3 shows that for the non-labor income elasticities the difference between the lowest and the highest elasticities is sizeable.

Table 1. Uncompensated wage elasticities

High: Woman aged 40, two children, non-labor income NOK 200,000, wage NOK70

Low: Woman aged 30, no children, non-labor income NOK 50,000, wage NOK 70

Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
High	0.620	0.670	0.465	0.649	0.660	0.599	1.309	1.374	1.093
Low	0.000	0.007	-0.043	0.163	0.152	0.222	0.163	0.160	0.178

Table 2. Compensated wage elasticities

High: Woman aged 40, two children, non-labor income NOK 50,000, wage NOK 300

Low: Woman aged 30, no children, non-labor income NOK 50,000, wage NOK 70

Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
High	0.419	0.331	0.765	0.267	0.259	0.292	0.687	0.590	1.057
Low	0.131	0.135	0.119	0.186	0.174	0.254	0.318	0.310	0.368

Table 3. Non-labor income elasticities

High: Woman aged 30, no children, non-labor income NOK 200,000, wage NOK 200

Low: Woman aged 30, two children, non-labor income NOK 100,000, wage NOK 70

Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
High	0.000	0.000	0.000	-0.010	-0.010	-0.013	-0.010	-0.010	-0.013
Low	-0.126	-0.137	-0.075	-0.149	-0.148	-0.146	-0.273	-0.284	-0.221

The uncompensated wage elasticities are higher at the intensive margin than at the extensive margin, with only two exceptions (out of 48 cases). The compensated wage elasticities are more equal at the extensive and the intensive margins, but with a weak tendency to be higher at the extensive margin. The non-labor income elasticities tend to be lowest (more negative) at the intensive margin.

The elasticities tend to decline with wage, in particular for women with two children. For example, for a woman aged 30 with two children and non-labor income NOK 200,000, the uncompensated elasticity of conditional expected hours (all sectors) declines from 0.634 (wage level NOK70) to 0.244 (wage level NOK 300) and the uncompensated elasticity of unconditional expected hours from 1.177 to 0.344. For the same woman and for the same wage levels, the decline in compensated elasticity is smaller: The compensated

elasticity of conditional expected hours drops from 0.410 to 0.266, and the compensated elasticity of unconditional expected hours drops from 1.093 to 0.635.

The participation elasticities are higher related to working in the private sector than in the public sector. The elasticities of conditional expected hours tend to be higher when the wage is low compared to when the wage is high. The reason is that when the wage is high the woman works long hours and hence the impact of a higher wage is lower compared to the case where the wage is low and initial hours are lower. This is particularly the case when the woman works in the private sector, where hours are less constrained.

If the woman has children the elasticities of conditional expected hours are higher compared to a childless woman. The reason is that the presence of small children reduces hours of work and hence the impact of a wage increase becomes stronger compared to a case for a childless women who works initially longer hours.

Recall that in the traditional deterministic microeconomic models the Slutsky equation holds: That is, the difference between the uncompensated and the compensated wage elasticities equals the non-labor income elasticity multiplied by the ratio of non-labor income to hours of work times the marginal wage. In random utility models, however, the Slutsky equation does not necessarily hold because choice behavior in this case is not determined by the usual marginal calculus. Table 4 below illustrates this feature. In particular, when non-labor income is equal to NOK 200,000 the Slutsky relation is very far from being met. See Appendix D for more detailed results.

Table 4. Comparison between uncompensated, compensated, and non-labor income elasticities. Woman aged 40, two children, wage NOK 70

Elasticity	Non-labor income NOK 50,000 Conditional mean hours			Non-labor income NOK 200,000 Conditional mean hours		
	All	Public	Private	All	Public	Private
Uncompensated	0.508	0.500	0.521	0.649	0.660	0.599
Compensated	0.566	0.557	0.571	0.412	0.376	0.487
Non-labor income	-0.145	-0.143	-0.146	-0.116	-0.118	-0.108

In Appendix F we show labor supply curves for a woman aged 30 and 40 with no children and two children. Notice that hours are along the vertical axis and wages along the horizontal axis. We show an uncompensated labor supply curve as well as a compensated. In Figures F.1 and F.2 we show the curves for the private and the public sector. The curves for the private sector are steeper and vary across shorter hours than the curves for the public sector. The uncompensated curves for both sectors taken together show how the expected hours, summed of the public and private sector using

uncompensated and unconditional choice probabilities, vary with wage rate. The same is done for the compensated labor supply curve, but using compensated choice probabilities, see Figures F.3 and F.4. The curves are upward sloping and slightly steeper at low wages and slightly steeper for women with children than without children. Note that taxes are accounted for when calculating the supply curves.

4.2. Numerical results using sample values

This section contains numerical results for compensated wage elasticities for the two-sector model using sample values. This allows us to show how the compensated elasticities vary across deciles of the (endogenous) income distribution.

The income decile limits are calculated from the expected household disposable income, using the probabilities (in B.1, Appendix B). For each possible hour we have calculated the mean of the probabilities based on 50x50 draws from the normal distribution to simulate the random error terms in the wage equations. By aggregating across all possible hours and sectors (including the not working option), we are then able to simulate incomes, and hence the income decile limits. Within each decile we then use eqs. B.4– B.13 to calculate the relative change in compensated probabilities for each level of hours of work and sector, and finally we calculate weighted average over hours, given the sector, using the compensated probabilities as weights. Again, we have to make draws from the normal distribution to simulate the error terms in the wage equations since these enter in (eqs. B.4–B.13). In order to calculate the elasticity of working, we take the weighted average over all deciles and sectors. To calculate the elasticity of working in a specific sector we do a similar calculation.

In Tables 5–7 we give the uncompensated and the compensated elasticities related to an overall wage increase. Table 5 gives the elasticities of working, and working in the two sectors; Table 6 gives the elasticities of conditional expected hours, conditional on working and working in specific sectors: and Table 7 gives the elasticities of the unconditional expected hours (which is the sum of the two above).

Table 5. Aggregated wage elasticities of the probability of working, and of working in the public or private sector, across deciles in the household income distribution

Sector	Uncompensated			Compensated		
	1 st decile	2 nd –9 th decile	10 th decile	1 st decile	2 nd –9 th decile	10 th decile
All sectors	0.2360	0.2805	0.2623	0.4170	0.4650	0.4752
Public	0.0790	0.0854	0.0811	0.2840	0.2670	0.2495
Private	0.3660	0.4880	0.4383	0.5232	0.6609	0.8180

Table 6. Aggregate compensated wage elasticities of conditional expected hours in the public or private sector, across deciles in the household income distribution

Sector	Uncompensated			Compensated		
	1 st decile	2 nd –9 th decile	10 th decile	1 st decile	2 nd –9 th decile	10 th decile
All sectors	0.3523	0.3529	0.3511	0.3069	0.2963	0.3256
Public	0.3700	0.3645	0.3655	0.3156	0.3052	0.324
Private	0.3373	0.3354	0.3337	0.3021	0.2903	0.3409

Table 7. Aggregate compensated wage elasticities of unconditional expected hours in the public or private sector, across deciles in the household income distribution

Sector	Uncompensated			Compensated		
	1 st decile	2 nd –9 th decile	10 th decile	1 st decile	2 nd –9 th decile	10 th decile
All sectors	0.5966	0.6433	0.6226	0.7238	0.7613	0.8009
Public	0.4519	0.4531	0.4495	0.6005	0.5772	0.5734
Private	0.7156	0.8399	0.7866	0.8253	0.9512	1.1590

A striking result is that both types of elasticities of conditional expected hours seem to be nearly the same across deciles and sectors. This is, however, a consequence of the distribution of the household characteristics of the sample, because we have seen in the previous section that the corresponding compensated and uncompensated elasticities differ within particular population groups and for different wage levels. For the private sector the compensated elasticities at the extensive margin tend to be higher than at the intensive margin. This is in accordance with the results reported in the previous section. The elasticities indicate that utility constant overall wage increases shift labor into the private sector and toward longer hours. We note that the uncompensated elasticities of conditional hours exceed the compensated elasticities in both sectors and all deciles, except for the 10th decile in the private sector. Thus these results also are inconsistent with the Slutsky equation.

Disregarding the stochastic components of the utility function

Recall from (3.2) that

$$(4.1a) \quad V(h, y) = \log(v(f(hw, y), h)\theta g(h)) + \eta(h)$$

For positive h and

$$(4.1b) \quad V(0, y) = \log(v(f(0, y), 0)) + \eta(0)$$

for $h = 0$, where the error terms are independent of the systematic terms and have the same distribution as the original error terms associated with the latent attributes of the jobs. In Tables 8–10 we report the elasticities that follow when we use only the mean value of the above utility function to simulate behavior. In this case it follows from the utility function above that the utility maximizing problem becomes equivalent to maximizing

the function $v(f(hw, y), h)\theta g(h)$. The motivation is to assess the significance of the stochastic elements of the utility function.

The elasticities in Tables 8 and 9 can be compared with the elasticities in Table 6 above. There are three important differences. First, both the uncompensated and compensated elasticities are numerically much higher than the elasticities that follow from the random utility model. Thus to conclude about wage and tax labor supply responses based on the deterministic parts of random utility models, might be quite misleading. Second, with the model represented by (4.1a,b), the elasticities in this case decline substantially from the first decile of the income distribution to the higher deciles. In the random utility model such a distinct pattern does not exist. Third, the compensated elasticities are significantly higher than the uncompensated, which implies also sizeable negative income elasticities, as shown in Table 10. With the random utility models these are numerically much lower, as indicated in Section 4.1.

Table 8. Uncompensated wage elasticities, given participation

Sector	Deciles		
	1 st decile	2 nd –9 th decile	10 th decile
Both sector	1.5907	0.8159	0.6887
Public	1.5035	0.7153	0.6204
Private	1.6585	0.9406	0.7570

Table 9. Compensated wage elasticities, given participation

Sector	Deciles		
	1 st decile	2 nd –9 th decile	10 th decile
Both sector	1.7546	0.9271	0.7601
Public	1.6800	0.8255	0.6912
Private	1.8125	1.0529	0.8290

Table 10. Income elasticities, given participation

Sector	Deciles		
	1 st decile	2 nd –9 th decile	10 th decile
Both sector	-0.6152	-0.2541	-0.1904
Public	-0.5633	-0.2143	-0.1714
Private	-0.6555	-0.3035	-0.2094

5. Conclusion

A major goal of this paper is to simplify, and in so-doing clarify, the analytic results obtained by Dagsvik and Karlström (2005) regarding compensated choice probabilities and the distribution of CV for random utility models. We have shown how the proofs of key results in Dagsvik and Karlström (2005) can be considerably simplified. A further aim has been to show how the apparatus of Dagsvik and Karlström (2005) can be applied to compute compensated choice probabilities for a particular labor supply model developed by Dagsvik and Strøm (2006). Finally, we have computed compensated and uncompensated elasticities for selected household groups, as well as for the sample used to estimate the model, based on the estimated labor supply model of Dagsvik and Strøm (2006). The numerical results demonstrate that the uncompensated and compensated wage elasticities may vary considerably between agents. The uncompensated wage elasticities are higher at the intensive margin than at the extensive margin, while the opposite tends to be the case for compensated elasticities. Both types of elasticity tend to decline by wage level. In the case of our particular sector model, we have also demonstrated that stochastic error terms play a major role for the numerical results.

While the Slutsky equation holds for the conventional labor supply model, this is no longer the case for models based on the random utility theory. In this paper we have shown empirically the extent to which the Slutsky equation is violated in our selected examples.

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Appendix A

For convenience we first prove the following elementary Lemma.

Lemma 1

$$\int_{-\infty}^{\infty} \exp(-e^{-x}(a+b))e^{-2x}b dx = \frac{b}{(a+b)^2}.$$

Proof of Lemma 1:

Integration by parts yields

$$b \int_{-\infty}^{\infty} \exp(-e^{-x}(a+b))e^{-2x} dx = \frac{b}{a+b} \Big|_{-\infty}^{\infty} e^{-x} \exp(-e^{-x}(a+b)) + \frac{b}{a+b} \int_{-\infty}^{\infty} \exp(-e^{-x}(a+b))e^{-x} dx$$

$$= \frac{b}{a+b} \int_{-\infty}^{\infty} \exp(-e^{-x}(a+b))e^{-x}dx = \frac{b}{(a+b)^2}.$$

Proof of Theorem 1:

We shall first derive the proof of $P^H(1,2)$ for the case with $m=3$. Assume that $y_2 < y_1$ and let $y \in [y_2, y_1]$. We have

$$\begin{aligned} \{J^0=1, J=2, Y=y\}1\{U_1^0=x\} &\Leftrightarrow \{\max(U_2^0, U_3^0) \leq U_1^0, \max(U_1(Y), U_3(Y)) \leq U_2(Y), Y=y\}1\{U_1^0=x\} \\ &= \{\max(U_2^0, U_3^0) \leq x, \max(U_1(y), U_3(y)) \leq U_2(y)=x\}1\{U_1^0=x\} \\ &= \{\max(U_3^0, U_3(y)) \leq x, U_2(y)=x, U_2^0 \leq x, U_1(y) \leq x\}1\{U_1^0=x\}. \end{aligned}$$

Since $y_2 \leq y \leq y_1$, it follows that $U_1(y) \leq U_1^0$ and $U_2(y) \geq U_2^0$. Hence, the event $\{U_2(y)=U_1^0=x\}$ implies that $\{U_1(y) \leq x, U_2^0 \leq x\}$. Accordingly

$$\{J^0=1, J=2, Y=y\} \Leftrightarrow \{\max(U_3^0, U_3(y)) \leq x, U_2(y)=x\}1\{U_1^0=x\}.$$

Thus, the corresponding probabilities are therefore given by

$$\begin{aligned} P(J^0=1, J=k, Y \in [y, y+\Delta y] | U_1^0=x) &= P(J^0=1, J=k, Y \in [y, y+\Delta y] | U_1^0=x) \\ &= P(\max(U_3^0, U_3(y)) \leq x, U_2(y) \leq x \leq U_2(y+\Delta y) | U_1^0=x) + o(\Delta y) \\ &= P(\max(U_3^0, U_3(y)) \leq x)P(U_2(y) \leq x < U_2(y+\Delta y)) + o(\Delta y) \\ &= P(\psi_3(y) + \varepsilon_3 \leq x)P(v_2(y) + \varepsilon_2 \leq x < v_2(y+\Delta y) + \varepsilon_2) + o(\Delta) \\ &= \exp(-e^{-x}\psi_3(y))(\exp(-e^{-x}v_2(y)) - \exp(-e^{-x}v_2(y+\Delta y))) + o(\Delta) \\ &= \exp(-e^{-x}\psi_3(y))\exp(-e^{-x}v_2(y))e^{-x}(v_2(y+\Delta y) - v_2(y)) + o(\Delta y). \end{aligned}$$

The last expression above implies that

$$\begin{aligned} (A.1) \quad P(\max(U_2^0, U_3^0) \leq x, \max(U_1(Y), U_3(Y)) \leq U_2(Y), Y \in dy) | U_1^0=x) \\ = \exp(-e^{-x}\psi_3(y))\exp(-e^{-x}v_2(y))e^{-x}v_2(dy). \end{aligned}$$

Thus, using (A.1) and Lemma 1, we obtain

$$\begin{aligned} (A.2) \quad P(J^0=1, J=2, Y \in [y, y+\Delta y]) &= EP(\max(U_2^0, U_3^0) \leq U_1^0, \max(U_1(Y), U_3(Y)) \leq U_2(Y), Y \in dy) | U_1^0) \\ &= \int_{-\infty}^{\infty} \exp(-e^{-x}\psi_3(y))\exp(-e^{-x}v_2(y))e^{-x} \exp(-e^{-x}v_1^0)e^{-x}v_1^0 dx \end{aligned}$$

$$= \frac{v_1^0 v_2(dy)}{\{v_1^0 + v_2(y) + \psi_3(y)\}^2}$$

from which Theorem 1 follows. In the general case with more than three alternatives we have

$$\begin{aligned} & P(J^0 = 1, J = 2, Y \in [y, y + \Delta y)) \\ &= P(\max_{r \neq 1} U_r^0 \leq x, \max_{r \neq 2} U_r(Y) \leq U_2(Y), Y \in (y, y + \Delta y) | U_1^0 = x) \\ &= P(\max_{r \neq 1} U_r^0 \leq x, \max_{r \neq 2} U_r(y) \leq x, U_2(y) \leq x < U_2(y + \Delta y) | U_1^0 = x) + o(\Delta) \\ &= P(\max_{r \neq \{1, 2\}} (U_r^0, U_r(y)) \leq x, U_2(y) \leq x < U_2(y + \Delta y) | U_1^0 = x) + o(\Delta) \\ &= P(\max_{r \neq \{1, 2\}} (\log \psi_r(y) + \varepsilon_r) \leq x, U_2(y) \leq x < U_2(y + \Delta y) | U_1^0 = x) + o(\Delta) \\ &= P(\max_{r \neq \{1, 2\}} (\log \psi_r(y) + \varepsilon_r) \leq x) P(U_2(y) \leq x < U_2(y + \Delta y) | U_1^0 = x) + o(\Delta). \end{aligned}$$

The rest of the proof in the general case is similar to the ternary case.

Q.E.D.

Proof of Corollary 1:

The result in (2.4) follows immediately from (2.3) because we have

$$P^H(j, k) = \int_{y_k}^{y_j} \varphi(j, k, dy).$$

Consider the result in (2.5). For this to happen it must be the case that $U_j^0 = U_j(Y)$, which implies that

$Y = y_j$. Furthermore, one must have

$$U_j^0 \geq \max_{r \neq j} (U_r^0, U_r(y_j)) = \max_{r \neq j} (\log \psi_r(y_j) + \varepsilon_j).$$

The corresponding probability is equal to

$$P^H(j, j) = \frac{v_j^0}{v_j^0 + \sum_{r=1, r \neq j}^m \psi_r(y_j)} = \frac{v_j^0}{\sum_{r=1}^m \psi_r(y_j)}.$$

The last equality follows from the fact that $v_j^0 = v_j(y_j)$.

Q.E.D.

Proof of Corollary 2:

Note that when $\{v_k(y)\}$ is continuously differentiable we have

$$\sum_k 1\{y > y_k\} v'_k(y) = \sum_k \psi'_k(y)$$

because $\psi_k(y)$ does not depend on y when $y \leq y_k$. Therefore, it follows that

$$\sum_{j,k} 1\{y_j \geq y \geq y_k\} v_j^0 v'_k(y) = \sum_j 1\{y_j > y\} v_j^0 \sum_k 1\{y > y_k\} v'_k(y) = \sum_j 1\{y_j > y\} v_j^0 \sum_k \psi_k(y).$$

Hence

$$\sum_{j,k} 1\{y_j \geq y \geq y_k\} \partial \varphi(j, k, y) / \partial = \frac{\sum_{j,k} 1\{y_j \geq y \geq y_k\} v_j^0 v_k'(y)}{\left\{ \sum_{r=1}^m \psi_r(y) \right\}^2} = \frac{\sum_j 1\{y_j \geq y\} v_j^0 \sum_{r=1}^m v_r'(y)}{\left\{ \sum_{r=1}^m \psi_r(y) \right\}^2}.$$

Q.E.D.

Proof of Corollary 3:

Recall that as in the proof above we have

$$\sum_k 1\{y > y_k\} v_k(dy) = \sum_k \psi_k(dy)$$

because $\psi_k(y)$ does not depend on y when $y \leq y_k$. Hence (2.2) implies that when $y \leq y_j$ then

$$P(J^0 = j, Y \in [y, y + dy)) = \sum_k 1\{y > y_k\} \varphi(j, k, dy) = \frac{v_j^0 \sum_k 1\{y > y_k\} v_k(dy)}{\left\{ \sum_{r=1}^m \psi_r(y) \right\}^2} = \frac{v_j^0 \sum_k \psi_k(dy)}{\left\{ \sum_{r=1}^m \psi_r(y) \right\}^2}.$$

By integrating the last expression with respect to y , the result of Corollary 3 follows.

Q.E.D.

Derivation of the distribution of $\max_{z \in B(h)} \varepsilon(z)$

By assumption

$$\begin{aligned} P\left(\max_{z \in B(h)} \varepsilon(z) - \log(\theta g(h)) \leq x\right) &= \prod_{z \in B(h)} P(\varepsilon(z) \leq x + \log(\theta g(h))) = \exp\left\{-\sum_{z \in B(h)} \exp(-x - \log(\theta g(h)))\right\} \\ &= \exp(-\theta g(h) \exp(-x - \log(\theta g(h)))) = \exp(-e^{-x}). \end{aligned}$$

The last equation follows from the fact that $\theta g(h)$ is the number of job alternatives in $B(h)$.

Q.E.D.

Appendix B

The two-sector discrete labor supply model

This section outlines the two-sector job choice model for married females. Here it is assumed that the female takes her husband's income as given. Let w_k denote the wage the female receives when working in sector k , $k = 1, 2$, and let $w = (w_1, w_2)$. The budget constraint when working in sector k , $k = 1, 2$, is given by

$$C_k = f(hw_k, I) \equiv hw_k - T(hw_k) + I,$$

where T is the tax function, h is hours of work and I is the sum of three income components. These three incomes are the after-tax wage income of the husband, the capital income (taxed at 28 percent) of the household,

and child allowances, which vary with the number of children up to the age of 18. Child allowances are not taxed. All details of the tax structure $T(\cdot)$ are taken into account in the estimation and simulation of the model. Let $\varphi_k(h)$ be the uncompensated probability of choosing a job in sector k with hours of work h (for a utility maximizing agent), and let D be the set of feasible hours (assumed to be the same across sectors). As in Section 3, it is demonstrated by Dagsvik and Strøm (2006) that

$$(B.1) \quad \varphi_k(h) = \frac{v(f(hw_k, I), h)g_k(h)\theta_k}{v(f(0, y), 0) + \sum_{r=1}^2 \sum_{x>0, x \in D} v(f(xw_r, I), x)g_r(x)\theta_r}$$

for $h > 0$, $k = 1, 2$, where $g_k(h)$ denotes the fraction of jobs with hours of work that are available in sector k (available to the agent). The term θ_k is a measure of the total number of jobs available to the female in sector k , $k = 1, 2$. For $h = 0$, $\varphi(0)$ is obtained from (B.1) by replacing the numerator with $v(f(0, y), 0)$.

The deterministic part of the utility function is specified as a Box-Cox transformation of consumption and leisure:

$$(B.2) \quad \log v(C, h) = \alpha_2 \left(\frac{[10^{-4}(C - C_0)]^{\alpha_1} - 1}{\alpha_1} \right) + \alpha_9 \left(\frac{[10^{-4}(C - C_0)]^{\alpha_1} - 1}{\alpha_1} \right) \left(\frac{(1 - h/3640)^{\alpha_3} - 1}{\alpha_3} \right) \\ + (\alpha_4 + \alpha_5 \log A + \alpha_6 (\log A)^2 + \alpha_7 X_{0,6} + \alpha_8 X_{7,17}) \left(\frac{(1 - h/3640)^{\alpha_3} - 1}{\alpha_3} \right).$$

Here C_0 is minimum consumption, A is age, and $X_{0,6}$, $X_{7,17}$ are the number of children below 6 and between 7 and 17 respectively. The alpha-s are unknown coefficients. If α_1 and α_3 are below 1, the deterministic part of the utility function is strictly concave. In order to make the paper self-contained, information about data, tax functions, and estimates is appended: see Tables C.1–C.4.

Consider now the calculation of the compensated choice probabilities in the context of a reform of the tax system or a change in wages. Here we shall assume that the opportunity measures $\theta_k g_k(h)$ remain unaffected by the reform. Let f^0 and f represent the initial and ex-post tax system and w_k^0 and w_k the initial and ex-post wage in sector k . Let $y_k(h)$ be defined by $v(f^0(hw_k^0, I), h) = v(f(hw_k, y_k(h)), h)$, for positive h and k , and let $y_0(0) = y(0)$, for $h = k = 0$. The function $y_k(h)$ is the ex-post non-labor income that makes the ex-ante deterministic part of utility equal to the corresponding ex-post part. Since the function v and the opportunity measures are unaffected by the reform it follows that the equation above is equivalent to

$$(B.3) \quad f^0(hw_k^0, I) = f(hw_k, y_k(h)).$$

Similarly to the definition in (3.2), let $P^H(j, h^0, k, h)$ denote the joint probability of being in sector j working h^0 hours ex-ante and working h hours in sector k ex post, given that the ex-ante and ex-post maximal utilities are the same. Here $j, k = 0, 1, 2$, where by sector 0 we understand the alternative not working. As in (3.3), it follows from Dagsvik and Karlström (2005) that

$$(B.4) \quad P^H(j, h^0, k, h) = 1\{y_j(h^0) > y_k(h)\} \int_{y_k(h)}^{y_j(h^0)} \frac{g_j(h^0)g_k(h)\theta_j\theta_k v(f^0(h^0 w_j^0, I), h^0) v(f(hw_k, dy), h)}{K(y)^2},$$

which is valid for positive hours and $k > 0, j > 0, (j, h^0) \neq (k, h)$. Due to (A.2), it follows that

$$(B.5) \quad K(y) = \max(v(f^0(0, I), 0), v(f(0, y), 0)) \\ + \sum_{r=1}^2 \sum_{x \in D} \max(g_r(x)\theta_r v(f^0(xw_r^0, I), x), g_r(x)\theta_r v(f(xw_r, y), x))$$

and

$$(B.6) \quad v(f(hw_k, dy), h) = \left(\alpha_2 + \alpha_9 \left(\frac{(1 - h/3640)^{\alpha_3} - 1}{\alpha_3} \right) \right) \left(10^{-4} (f(hw_k, y) - C_o) \right)^{\alpha_1 - 1} v(f(hw_k, y), h) dy.$$

The cases with not working alternatives we obtain that:

$$(B.7) \quad P^H(j, h^0, 0, 0) = 1\{y_j(h^0) > y_0(0)\} \int_{y(0)}^{y_j(h^0)} \frac{g_j(h^0)\theta_j v(f^0(h^0 w_k^0, I), h^0) v(f(0, dy), 0)}{K(y)^2},$$

$$(B.8) \quad P^H(0, 0, k, h) = 1\{y_0(0) > y_k(h)\} \int_{y_k(h)}^{y_0(0)} \frac{g_k(h)\theta_k v(f^0(0, I), 0) v(f(hw_k, dy))}{K(y)^2},$$

$$(B.9) \quad P^H(j, h, j, h) = \frac{g_j(h)\theta_j v(f^0(hw_k^0, I), h)}{K(y_j(h))},$$

for positive h and j , and

$$(B.10) \quad P^H(0, 0, 0, 0) = \frac{v(f^0(0, I), 0)}{K(y(0))}$$

for $j = h = 0$.

Let $\varphi_k^H(h)$ be the probability that the agent chooses (k, h) ex-post, given that utility is the same as before the reform. Then for $h > 0$,

$$(B.11) \quad \varphi_k^H(h) = \sum_{j \neq k} \sum_{x \in D \setminus \{h\}} P^H(j, x, k, h) + \sum_{x \in D \setminus \{h\}} P^H(k, x, k, h) + \sum_{j \neq k} P^H(j, h, k, h) + P^H(k, h, k, h),$$

for $h, k > 0$, and

$$(B.12) \quad \varphi^H(0) = \sum_{x \in D} \sum_{j > 0} P^H(j, x, 0, 0) + P^H(0, 0, 0, 0).$$

The compensated and the relative change for all (h, k) is given by

$$(B.13) \quad \varphi_k^H(h) - \varphi_k^0(h) \quad \text{and} \quad \frac{\varphi_k^H(h) - \varphi_k^0(h)}{\varphi_k^0(h)}.$$

Appendix C

Data, tax functions, and estimates

Data on the labor supply of married women in Norway used in this study consist of a merged sample of the “Survey of Income and Wealth, 1994” and the “Level of living conditions, 1995” (Statistics Norway, 1994 and 1995 respectively). Data cover married couples as well as cohabiting couples with common children. The ages of the spouses range from 25 to 64. None of the spouses is self-employed and none of them is on disability or other types of benefits. A person is classified as a wage worker if their income from wage work is higher than their income from self-employment. All taxes paid are observed and in the assessment of disposable income, at hours not observed, all details of the tax system are accounted for. Hours of work are calculated as the sum of hours of the main job as well as those of any subsidiary jobs. A large majority of the women have only one job.

Wage rates above NOK 350 or below NOK 40⁶ are not utilized when estimating the wage equations. The wage rates are computed as the ratio of annual wage income to hours worked. When computing annual wage income, we take into account the fact that some women have multiple jobs. The size of the sample used in estimating the labor supply model is 810. Descriptions of variables and summary statistics are given in Table C.1.

Table C.1. Descriptive statistics, number of observations = 810 (values in NOK, 1994)

Variable	Mean	Std Dev.	Min.	Max.
Number of children (age 0–2)	0.23	0.45	0.00	2.00
Number of children (age 0–6)	0.54	0.77	0.00	3.00
Number of children (age 3–6)	0.30	0.56	0.00	3.00
Number of children (age 7–17)	0.66	0.85	0.00	4.00
Age in year (men)	42.80	9.17	25.00	66.00
Education in year (men)	12.05	2.49	9.00	19.00
Age in year (women)	40.07	9.04	25.00	64.00
Education in year (women)	11.61	2.15	9.00	17.00
Sector (1 = Public, 2 = Private)	1.34	0.61	0.00	2.00
Work experience (woman age – woman education in years)	22.45	9.63	2.00	49.00
Capital income (child allowances included)	32306.71	42378.48	0.00	568403.00
Child allowances	13094.37	12154.01	0.00	60084.00
Women wage income per year	149751.97	83060.53	0.00	581693.00
Men wage income per year	274372.89	106239.67	17312.00	1184861.00
Woman hourly wage in public sector	89.36	12.09	64.88	132.34
Woman hourly wage in private sector	109.77	13.68	80.14	156.44

⁶ As of June 2013, 1 USD = NOK 5.80.

Table C.2. Estimates of the model parameters

Variables	Parameters	Estimates	t-values
Preferences			
<i>Consumption</i>			
Exponent	α_1	0.64	7.6
Scale 10^{-4}	α_2	1.77	4.2
Subsistence level C_0 in NOK per year		60 000	
<i>Leisure</i>			
Exponent	α_3	-0.53	-2.1
Constant	α_4	111.66	3.2
Log age	α_5	-63.61	-3.2
$(\log \text{age})^2$	α_6	9.2	3.3
# children 0–6	α_7	1.27	4.0
# children 7–17	α_8	0.97	4.1
<i>Consumption and Leisure, interaction</i>			
Subsistence level of leisure in hours per year	α_9	-0.12	-2.7
		5120	
The parameters θ_1 and θ_2; $\log \theta_j = f_{j1} + f_{j2}S$			
Constant, public sector (sector 1)	f_{11}	-4.2	-4.7
Constant, private sector (sector 2)	f_{21}	1.14	1.0
Education, public sector (sector 1)	f_{12}	0.22	2.9
Education, private sector (sector 2)	f_{22}	-0.34	-3.3
Opportunity density of offered hours, $g_{k2}(h)$, $k=1,2$			
Full-time peak, public sector (sector 1)*	$\log(g_{12}(h_{\text{Full}}) / g_{12}(h_0))$	1.58	11.8
Full-time peak, private sector (sector 2)	$\log(g_{22}(h_{\text{Full}}) / g_{22}(h_0))$	1.06	7.4
Part-time peak, public Sector	$\log(g_{12}(h_{\text{Par}}) / g_{12}(h_0))$	0.68	4.4
Part-time peak, private Sector	$\log(g_{22}(h_{\text{Par}}) / g_{22}(h_0))$	0.8	5.2
# observations		824	
Log likelihood		-1760.9	

* The notation h_0 refers to an arbitrary level of hours of work different from full-time and part-time hours.

The deterministic part of the utility function, i.e. the numerator in the choice probabilities, given participation, is quasi-concave for all individuals in the sample.

Table C.3. Tax function in 1994 for a married non-working woman whose husband is working, 1994

Male income, Y_{male}	Tax T
0–41,907	0
41,907–140,500	$0.302Y_{\text{male}} - 12656$
140,500–252,000	$0.358Y_{\text{male}} - 20524$
252,000–263,000	$0.453Y_{\text{male}} - 44464$
263,000–	$0.495Y_{\text{male}} - 55510$

Table C. 4. Tax function in 1994 for a married working woman or man, NOK 1994

Wage income, Y	Tax T
0–20,954	0
20,954–140,500	$0.302Y - 6328$
140,500–208,000	$0.358Y - 14196$
208,000–236,500	$0.453Y - 33956$
236,500–	$0.495Y - 43889$

Appendix D

Elasticities

Table D.1. Married woman aged 30, no children

Wage NOK 70

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	0.007	-0.043	0.163	0.152	0.222	0.163	0.160	0.178
	Compensated	0.131	0.135	0.111	0.186	0.174	0.254	0.318	0.310	0.365
	Income	0.000	-0.007	0.040	-0.060	-0.057	-0.081	-0.060	-0.064	-0.041
100,000	Uncompensated	0.029	0.057	-0.089	0.269	0.254	0.352	0.299	0.306	0.259
	Compensated	0.183	0.205	0.061	0.254	0.239	0.335	0.438	0.445	0.397
	Income	-0.022	-0.031	0.028	-0.080	-0.076	-0.101	-0.102	-0.107	-0.073
200,000	Uncompensated	0.119	0.149	-0.038	0.355	0.338	0.441	0.479	0.493	0.401
	Compensated	0.276	0.324	0.015	0.319	0.305	0.389	0.596	0.629	0.404
	Income	-0.046	-0.054	-0.009	-0.062	-0.060	-0.074	-0.109	-0.114	-0.083

Wage NOK 200

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	-0.046	0.245	0.192	0.190	0.194	0.192	0.142	0.445
	Compensated	0.295	0.237	0.588	0.224	0.223	0.222	0.519	0.461	0.810
	Income	0.000	0.001	-0.007	-0.008	-0.008	-0.009	-0.008	-0.007	-0.017
100,000	Uncompensated	0.000	-0.043	0.231	0.207	0.203	0.220	0.207	0.158	0.457
	Compensated	0.272	0.219	0.546	0.221	0.219	0.225	0.493	0.438	0.771
	Income	0.000	0.001	-0.007	-0.012	-0.011	-0.014	-0.012	-0.010	-0.021
200,000	Uncompensated	0.000	-0.039	0.202	0.216	0.211	0.239	0.216	0.172	0.453
	Compensated	0.240	0.194	0.481	0.215	0.212	0.227	0.456	0.406	0.709
	Income	0.000	0.000	0.000	-0.010	-0.010	-0.013	-0.010	-0.010	-0.013

Wage NOK 300

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	-0.049	0.229	0.178	0.184	0.147	0.178	0.134	0.380
	Compensated	0.347	0.297	0.574	0.169	0.175	0.138	0.516	0.472	0.713
	Income	0.000	0.001	-0.006	-0.004	-0.005	-0.004	-0.004	-0.003	-0.010
100,000	Uncompensated	0.000	-0.051	0.240	0.186	0.192	0.155	0.186	0.140	0.399
	Compensated	0.338	0.286	0.570	0.173	0.179	0.142	0.511	0.465	0.713
	Income	0.000	0.002	-0.009	-0.007	-0.007	-0.006	-0.007	-0.005	-0.015
200,000	Uncompensated	0.000	-0.052	0.249	0.195	0.200	0.165	0.195	0.147	0.418
	Compensated	0.315	0.263	0.549	0.175	0.181	0.146	0.490	0.444	0.695
	Income	0.000	0.001	-0.008	-0.007	-0.007	-0.006	-0.007	-0.005	-0.015

Table D.2. Woman aged 30, two children
Wage NOK 70

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.047	0.096	-0.190	0.462	0.451	0.498	0.511	0.551	0.299
	Compensated	0.327	0.394	-0.012	0.494	0.481	0.536	0.821	0.876	0.523
	Income	-0.046	-0.059	0.014	-0.131	-0.128	-0.137	-0.177	-0.187	-0.122
100,000	Uncompensated	0.254	0.304	0.021	0.593	0.589	0.595	0.862	0.911	0.618
	Compensated	0.581	0.704	-0.017	0.667	0.675	0.569	1.249	1.379	0.551
	Income	-0.126	-0.137	-0.075	-0.149	-0.148	-0.146	-0.273	-0.284	-0.221
200,000	Uncompensated	0.511	0.554	0.319	0.634	0.637	0.603	1.177	1.227	0.942
	Compensated	0.682	0.830	-0.027	0.410	0.382	0.508	1.093	1.212	0.480
	Income	-0.167	-0.173	-0.141	-0.111	-0.112	-0.105	-0.277	-0.283	-0.246

Wage NOK 200

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	0.001	-0.006	0.201	0.188	0.277	0.201	0.189	0.271
	Compensated	0.246	0.232	0.325	0.212	0.200	0.287	0.458	0.432	0.613
	Income	0.000	-0.001	0.009	-0.014	-0.013	-0.020	-0.014	-0.015	-0.011
100,000	Uncompensated	0.000	0.008	-0.048	0.231	0.216	0.319	0.231	0.222	0.270
	Compensated	0.247	0.241	0.281	0.227	0.213	0.310	0.475	0.455	0.591
	Income	0.000	-0.003	0.018	-0.025	-0.020	-0.034	-0.025	-0.027	-0.016
200,000	Uncompensated	0.001	0.019	-0.101	0.278	0.261	0.381	0.280	0.281	0.276
	Compensated	0.249	0.253	0.221	0.254	0.238	0.347	0.503	0.492	0.568
	Income	0.000	-0.006	0.027	-0.035	-0.032	-0.046	-0.035	-0.038	-0.019

Wage NOK 300

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	-0.053	0.296	0.229	0.223	0.252	0.229	0.169	0.556
	Compensated	0.440	0.367	0.826	0.273	0.270	0.284	0.714	0.637	1.110
	Income	0.000	0.001	-0.005	-0.006	-0.006	-0.007	-0.006	-0.005	-0.013
100,000	Uncompensated	0.000	-0.050	0.282	0.234	0.228	0.264	0.234	0.176	0.553
	Compensated	0.413	0.345	0.778	0.268	0.263	0.284	0.681	0.608	1.063
	Income	0.000	0.001	-0.006	-0.010	-0.009	-0.012	-0.010	-0.008	-0.019
200,000	Uncompensated	0.000	-0.045	0.253	0.244	0.235	0.284	0.244	0.189	0.544
	Compensated	0.375	0.315	0.702	0.260	0.253	0.285	0.635	0.568	0.988
	Income	0.000	0.000	-0.002	-0.013	-0.012	-0.016	-0.013	-0.011	-0.019

Table D.3. Married woman aged 40, no children

Wage NOK 70

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	0.019	-0.078	0.187	0.172	0.249	0.187	0.192	0.168
	Compensated	0.148	0.164	0.083	0.211	0.195	0.283	0.360	0.359	0.366
	Income	0.000	-0.012	0.050	-0.076	-0.070	-0.098	-0.076	-0.082	-0.048
100,000	Uncompensated	0.045	0.085	-0.107	0.319	0.298	0.399	0.366	0.384	0.287
	Compensated	0.225	0.271	0.042	0.303	0.282	0.380	0.582	0.553	0.422
	Income	-0.034	-0.049	0.022	-0.100	-0.094	-0.120	-0.134	-0.143	-0.098
200,000	Uncompensated	0.177	0.226	-0.004	0.417	0.395	0.490	0.601	0.631	0.486
	Compensated	0.364	0.458	0.000	0.389	0.372	0.433	0.753	0.831	0.433
	Income	-0.069	-0.080	-0.030	-0.078	-0.075	-0.088	-0.147	-0.155	-0.119

Wage NOK 200

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	-0.053	0.209	0.199	0.192	0.218	0.199	0.138	0.432
	Compensated	0.291	0.220	0.563	0.229	0.224	0.240	0.521	0.445	0.804
	Income	0.000	0.001	-0.005	-0.009	-0.009	-0.011	-0.009	-0.007	-0.016
100,000	Uncompensated	0.000	-0.049	0.193	0.209	0.200	0.236	0.209	0.150	0.434
	Compensated	0.270	0.205	0.517	0.226	0.220	0.244	0.498	0.426	0.762
	Income	0.000	0.001	-0.004	-0.013	-0.013	-0.016	-0.013	-0.011	-0.021
200,000	Uncompensated	0.000	-0.041	0.165	0.223	0.212	0.260	0.223	0.170	0.430
	Compensated	0.241	0.187	0.450	0.223	0.214	0.249	0.464	0.402	0.699
	Income	0.000	0.000	0.003	-0.014	-0.013	-0.018	-0.014	-0.013	-0.014

Wage NOK 300

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	-0.070	0.258	0.200	0.207	0.172	0.200	0.138	0.427
	Compensated	0.393	0.315	0.650	0.198	0.206	0.167	0.591	0.521	0.817
	Income	0.000	0.002	-0.007	-0.005	-0.005	-0.004	-0.005	-0.003	-0.011
100,000	Uncompensated	0.000	-0.074	0.259	0.208	0.214	0.181	0.208	0.138	0.446
	Compensated	0.379	0.301	0.640	0.201	0.208	0.171	0.581	0.510	0.812
	Income	0.000	0.002	-0.010	-0.008	-0.008	-0.007	-0.008	-0.005	-0.017
200,000	Uncompensated	0.000	-0.075	0.265	0.217	0.222	0.193	0.217	0.145	0.463
	Compensated	0.351	0.276	0.610	0.202	0.209	0.175	0.554	0.483	0.786
	Income	0.000	0.002	-0.009	-0.008	-0.008	-0.007	-0.008	-0.005	-0.016

Table D.4. Married woman aged 40, two children

Wage NOK 70

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.065	0.131	-0.153	0.508	0.500	0.521	0.578	0.638	0.360
	Compensated	0.371	0.474	0.011	0.566	0.557	0.571	0.938	1.031	0.582
	Income	-0.065	-0.082	-0.010	-0.145	-0.143	-0.146	-0.210	-0.224	-0.156
100,000	Uncompensated	0.331	0.393	0.130	0.629	0.631	0.605	0.981	1.049	0.744
	Compensated	0.714	0.920	0.006	0.784	0.820	0.576	1.498	1.740	0.582
	Income	-0.163	-0.176	-0.122	-0.158	-0.160	-0.151	-0.320	-0.334	-0.271
200,000	Uncompensated	0.620	0.670	0.465	0.649	0.660	0.599	1.309	1.374	1.093
	Compensated	0.821	1.063	-0.008	0.412	0.376	0.487	1.233	1.440	0.478
	Income	-0.203	-0.209	-0.183	-0.116	-0.118	-0.108	-0.317	-0.325	-0.290

Wage NOK 200

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	0.014	-0.062	0.218	0.199	0.297	0.218	0.214	0.233
	Compensated	0.257	0.253	0.273	0.223	0.204	0.301	0.480	0.458	0.575
	Income	0.000	-0.003	0.013	-0.018	-0.016	-0.024	-0.018	-0.019	-0.011
100,000	Uncompensated	0.000	0.026	-0.105	0.257	0.235	0.348	0.258	0.265	0.239
	Compensated	0.265	0.272	0.234	0.245	0.225	0.332	0.511	0.497	0.567
	Income	0.000	-0.005	0.023	-0.031	-0.028	-0.041	-0.031	-0.034	-0.018
200,000	Uncompensated	0.002	0.042	-0.156	0.316	0.291	0.419	0.319	0.335	0.255
	Compensated	0.273	0.295	0.181	0.284	0.261	0.380	0.557	0.557	0.561
	Income	-0.001	-0.009	0.032	-0.044	-0.040	-0.056	-0.045	-0.050	-0.024

Wage NOK 300

Income	Elasticity	Probability of working			Conditional mean hours			Unconditional mean hours		
		All	Public	Private	All	Public	Private	All	Public	Private
50,000	Uncompensated	0.000	-0.059	0.244	0.221	0.211	0.255	0.221	0.149	0.506
	Compensated	0.419	0.331	0.765	0.267	0.259	0.292	0.687	0.590	1.057
	Income	0.000	0.000	-0.003	-0.006	-0.006	-0.008	-0.006	-0.005	-0.012
100,000	Uncompensated	0.000	-0.054	0.225	0.228	0.216	0.269	0.228	0.160	0.501
	Compensated	0.395	0.315	0.715	0.261	0.252	0.292	0.657	0.567	1.008
	Income	0.000	0.000	-0.003	-0.011	-0.010	-0.014	-0.011	-0.009	-0.018
200,000	Uncompensated	0.000	-0.046	0.190	0.243	0.228	0.295	0.243	0.180	0.491
	Compensated	0.363	0.294	0.640	0.255	0.244	0.295	0.619	0.538	0.935
	Income	0.000	0.000	0.001	-0.015	-0.014	-0.020	-0.015	-0.014	-0.018

Appendix E

Uncompensated probabilities, conditional, and unconditional mean hours

Table E.1 Married woman aged 30, no children

Wage NOK 70

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.8504	0.1496	1907	1912	1878	1907	1626	281
100,000	0.9936	0.8399	0.1537	1810	1821	1755	1799	1529	270
200,000	0.9691	0.8140	0.1551	1711	1725	1637	1658	1404	254

Wage NOK 200

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.8410	0.1590	2185	2176	2230	2185	1830	355
100,000	1.0000	0.8419	0.1581	2168	2160	2211	2168	1819	350
200,000	1.0000	0.8424	0.1576	2149	2142	2187	2149	1804	345

Wage NOK 300

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.8228	0.1772	2377	2369	2415	2377	1949	428
100,000	1.0000	0.8238	0.1762	2367	2359	2406	2367	1943	424
200,000	1.0000	0.8251	0.1749	2354	2346	2395	2354	1935	419

Table E.2 Married woman aged 30, two children

Wage NOK 70

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	0.9894	0.8200	0.1693	1496	1515	1403	1480	1243	238
100,000	0.9331	0.7664	0.1667	1348	1366	1261	1257	1047	210
200,000	0.8352	0.6814	0.1537	1218	1235	1144	1017	842	176

Wage NOK 200

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.8540	0.1460	1956	1958	1945	1956	1672	284
100,000	1.0000	0.8525	0.1474	1929	1933	1907	1929	1648	281
200,000	0.9997	0.8498	0.1499	1887	1894	1851	1887	1609	278

Wage NOK 300

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.8470	0.1530	2132	2124	2173	2132	1799	332
100,000	1.0000	0.8477	0.1523	2119	2112	2157	2119	1791	329
200,000	1.0000	0.8483	0.1517	2101	2095	2135	2101	1777	324

Table E.3. Married woman aged 40, no children

Wage NOK 70

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.8020	0.1980	1855	1865	1812	1855	1496	359
100,000	0.9898	0.7858	0.2040	1739	1756	1671	1721	1380	341
200,000	0.9529	0.7490	0.2039	1621	1643	1540	1545	1231	314

Wage NOK 200

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.7967	0.2033	2141	2131	2179	2141	1698	443
100,000	1.0000	0.7975	0.2025	2123	2115	2157	2123	1686	437
200,000	1.0000	0.7977	0.2023	2101	2094	2128	2101	1670	431

Wage NOK 300

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.7745	0.2255	2337	2326	2378	2337	1801	536
100,000	1.0000	0.7759	0.2241	2326	2314	2367	2326	1795	531
200,000	1.0000	0.7776	0.2224	2311	2299	2354	2311	1788	524

Table E.4. Married woman aged 40, two children
Wage NOK 70

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	0.9848	0.7594	0.2254	1416	1440	1332	1394	1094	300
100,000	0.9099	0.6937	0.2161	1263	1286	1191	1149	892	257
200,000	0.7916	0.5990	0.1926	1136	1154	1078	899	691	208

Wage NOK 200

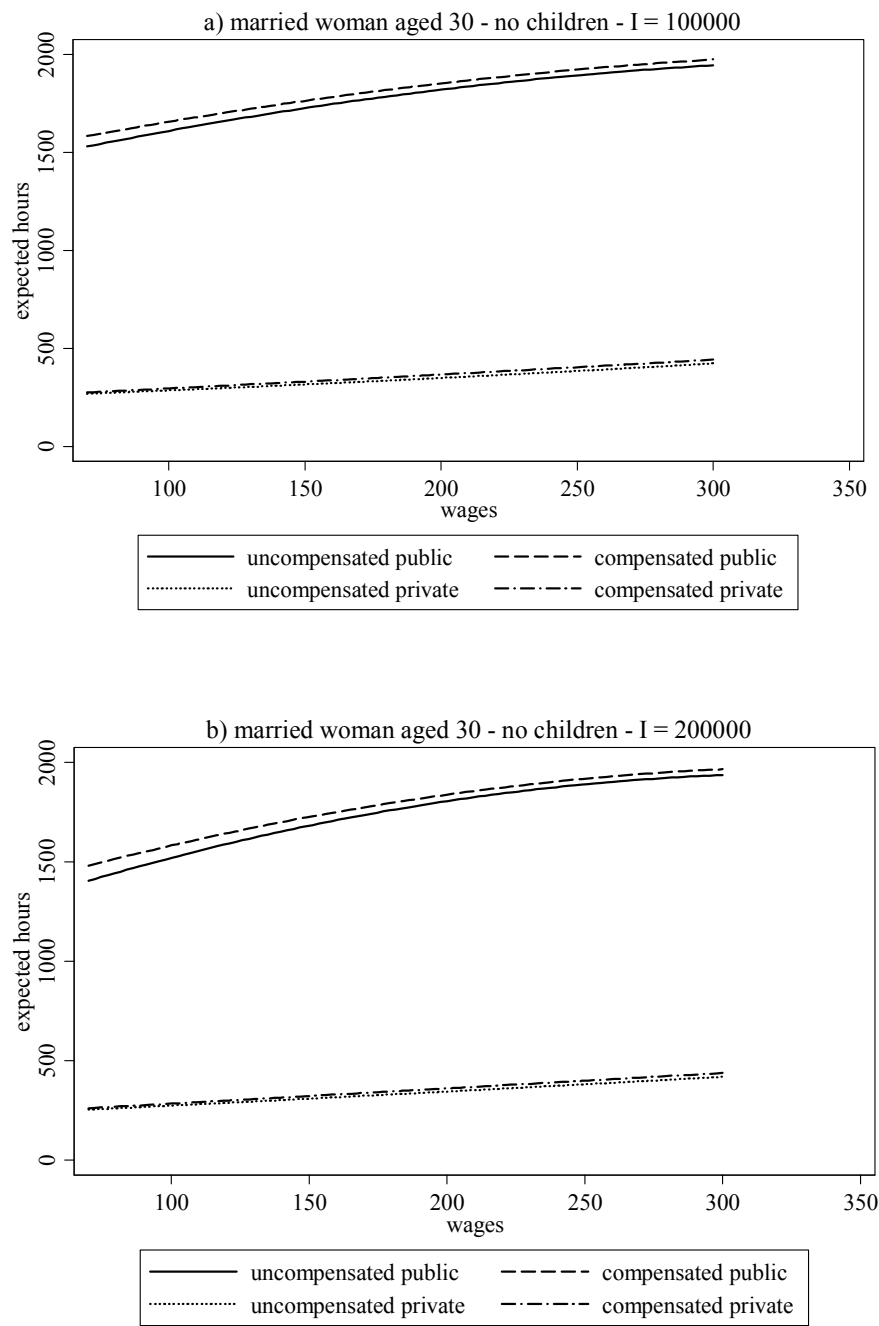
Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.8075	0.1925	1919	1925	1896	1919	1554	365
100,000	1.0000	0.8051	0.1949	1887	1895	1853	1887	1526	361
200,000	0.9995	0.8006	0.1989	1836	1848	1788	1835	1480	356

Wage NOK 300

Income	Probability of working			Conditional mean hours			Unconditional mean hours		
	All	Public	Private	All	Public	Private	All	Public	Private
50,000	1.0000	0.8034	0.1966	2096	2088	2129	2096	1678	418
100,000	1.0000	0.8040	0.1960	2083	2076	2112	2083	1669	414
200,000	1.0000	0.8043	0.1957	2063	2058	2086	2063	1655	408

Appendix F. Labor supply curves

Figure F.1. a-d. Uncompensated and compensated unconditional expected hours, public and private sector. Woman aged 30.



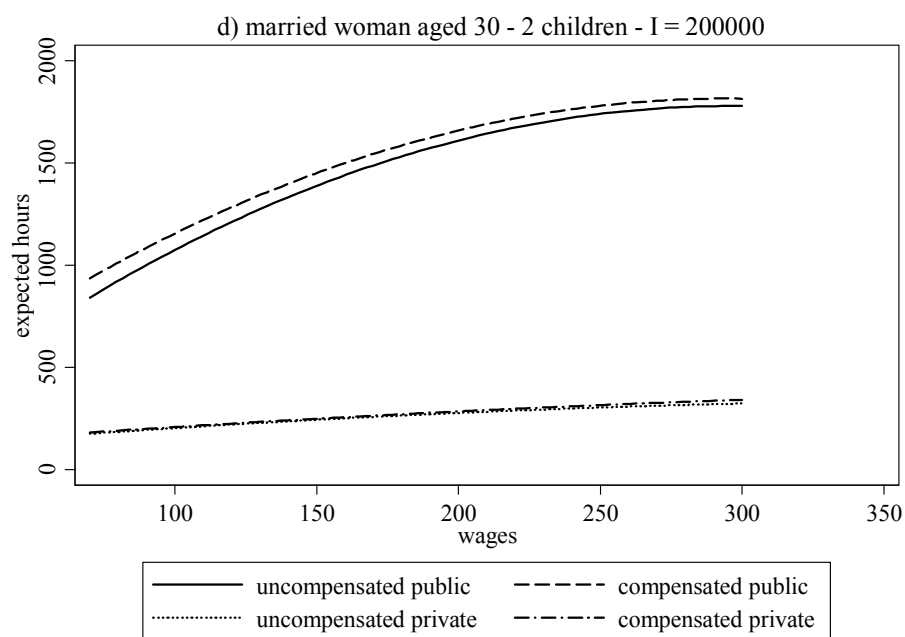
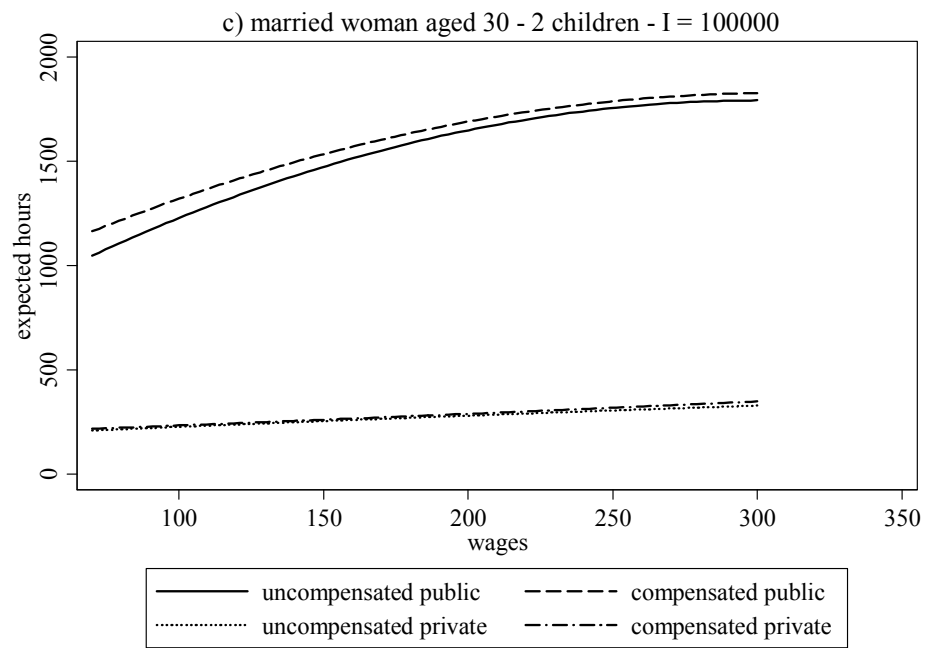
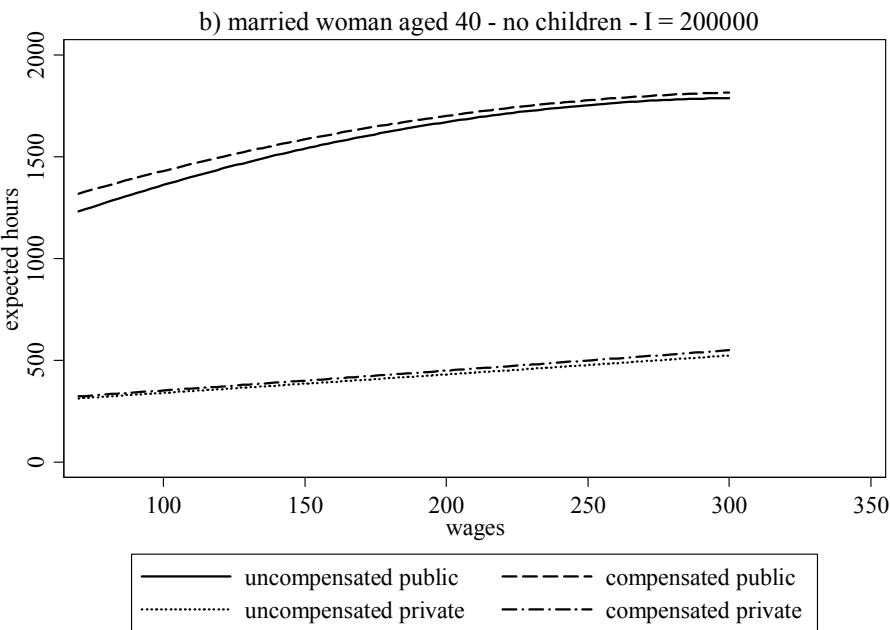


Figure F. 2.a-d. Uncompensated and compensated unconditional expected hours, public and private sector: woman aged 40.



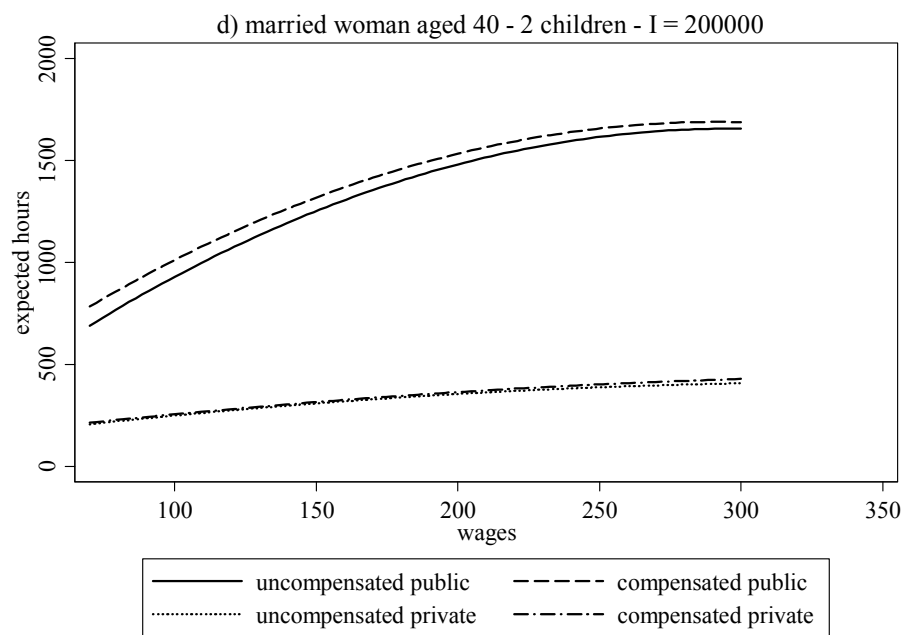
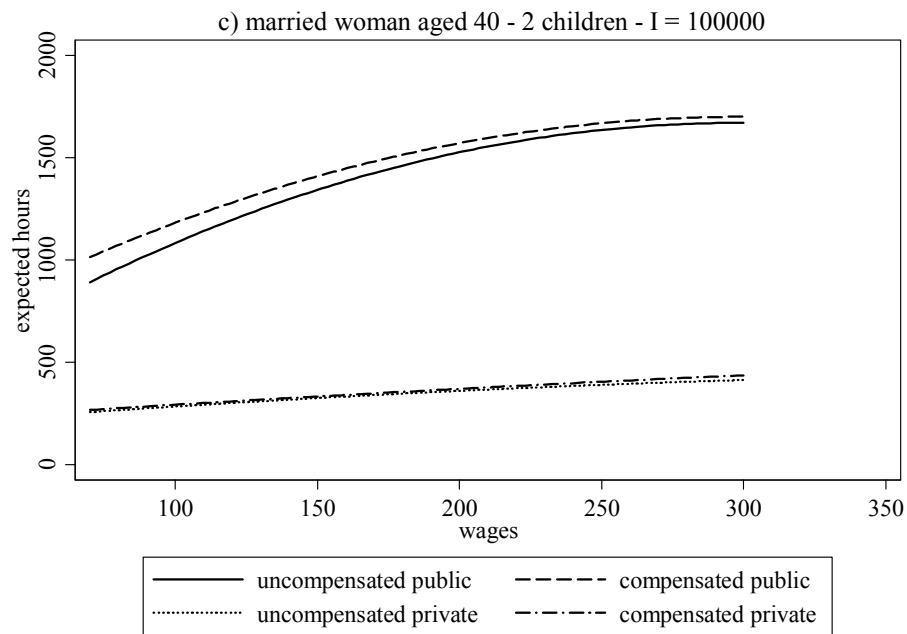
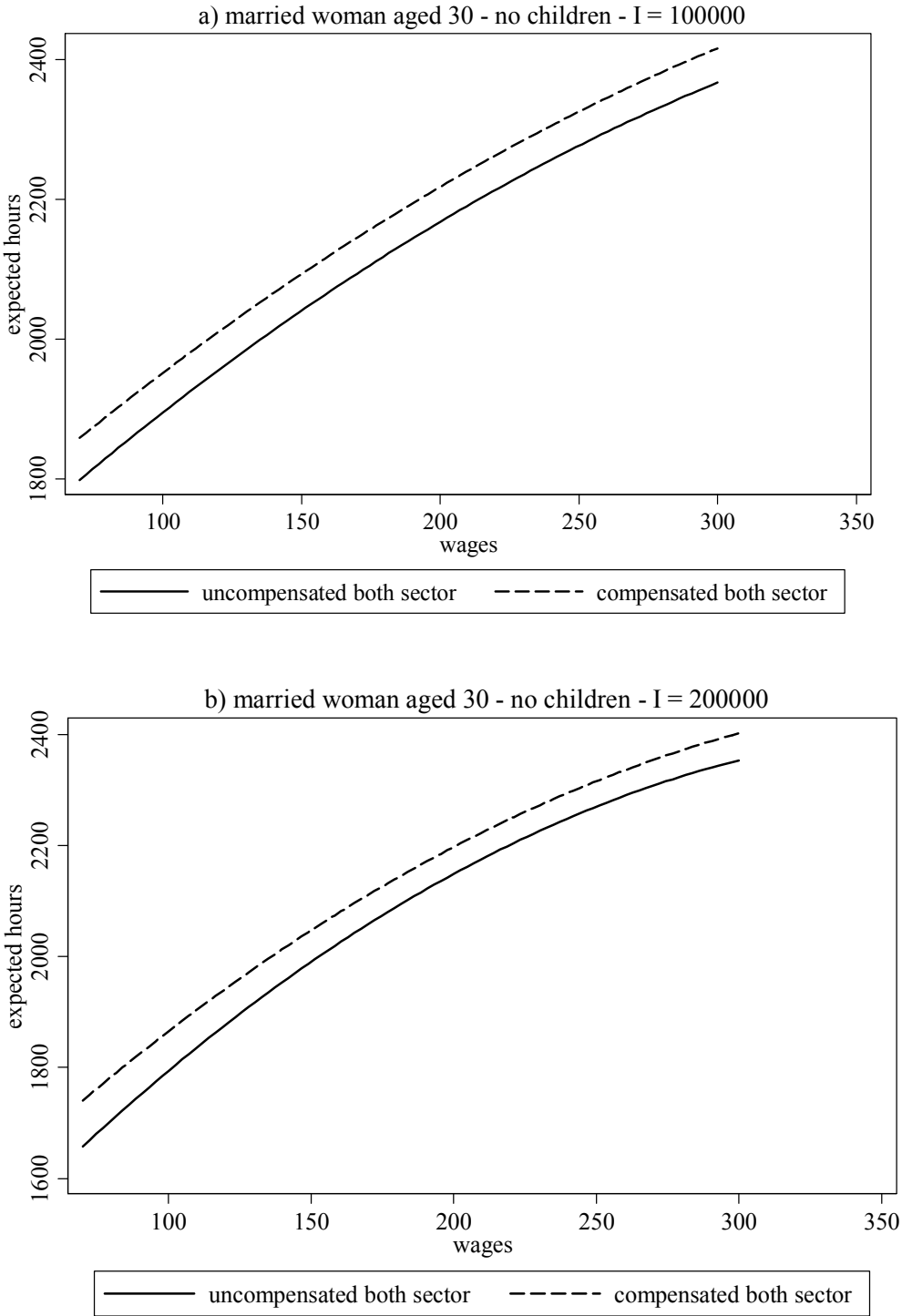
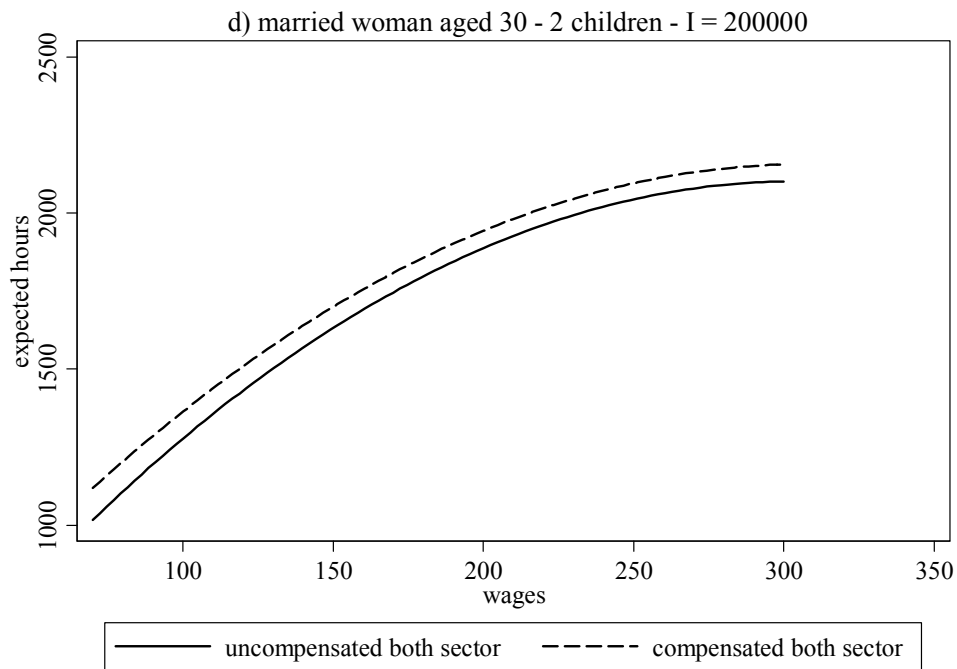


Figure F.3. a-d. Uncompensated and compensated unconditional expected hours, both sectors: woman aged 30.





**Figure F.4.a-d. Uncompensated and compensated unconditional expected hours, both sector:
woman aged 40.**

