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Family Structure and the Education Gender Gap: Evidence from Italian Provinces

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Abstract

We investigate the determinants of the education gender gap in Italy in historical perspective with a focus on the influence of family structure. We capture the latter with two indicators: residential habits (nuclear vs. complex families) and inheritance rules (partition vs. primogeniture). After controlling for economic, institutional, religious, and cultural factors, we find that family structure is a driver of the education gender gap, with a higher female to male enrollment rate ratio in upper primary schools being associated with nuclear residential habits and equal partition of inheritance. The effect tends to persist through the present day.


Keywords: education gender gap, Italian unification, family types, inheritance, institutions, religion, convergence.

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1. Introduction

Over a century and a half after the Unification of the country in 1861, in Italy the education gender gap has been gradually closing, and even reversing in recent years. This paper investigates its determinants by exploring the comparative effect of economic, institutional, religious, and cultural factors in historical perspective. The main focus of our investigation is the potential influence of family structure on gendered human capital accumulation. Despite the fact that Italy – within cross-country investigations – is often treated as homogeneous in this respect, in fact it does exhibit significant variations in its family structure along several dimensions.

The industrialization process taking place in the 1950s and 1960s has profoundly affected the Italian family structure by provoking a shift of residential habits toward the nuclear model. However, the nuclear model, as opposed to more complex ones, was already diffused even during the nineteenth century, especially in the South. Likewise, partible inheritance rules have long dominated a large portion of the peninsula. At the same time, however, a patriarchal and authoritarian attitude was also prevalent within Italian families, according to a rigid hierarchy that emphasized gender roles (Barbagli, 1987). Despite the fact that family structure, in Italy and elsewhere, not only went through a time evolution, but also responded to different rules in rural vs. urban contexts and also across different social classes, it is possible to identify a relatively stable geographical differentiation of family models by the time the country is unified. To classify family structure, we monitor two specific dimensions, reflecting residential habits and rules of inheritance, respectively. Therefore, we distinguish between nuclear vs. complex residential habits and between equal partition vs. primogeniture. Nuclear residential habits reflect a more liberal, rather than authoritarian, attitude which should favour a more advantageous position of women within the family and in society. Partible inheritance is associated with more equality among children including females, if compared to alternative customs such as male primogeniture or discretion, which typically favour male heirs. We also evaluate the potential additional role of female age at marriage, since age at marriage is another important indicator of the degree of subordination of women with respect to men and therefore of gender inequality. Moreover, we combine these three indicators – residential habits, inheritance rules, and female age at marriage – to obtain an extended version, adapted to Italy, of the classification of family structure due to Todd (1990). While Todd only considers residential habits and inheritance rules, we also evaluate the combined potential role of female age at marriage, following the attention paid by Hajnal (1965, 1982), and by Barbagli (1987) and Del Panta et al. (1996) with specific focus on Italy, to this additional dimension.

Our analysis of the effect of family structure on the education gender gap is based on a comprehensive newly-assembled database reporting primary school enrollment, by gender, across 69 provinces over a twenty-year sub-sample covering the 1861-1901 period which captures the crucial initial 40 years of the country. At this critical juncture for the history of the country, we still observe a sizeable gender gap even in compulsory primary school. We complete our database with information on relative contemporaneous schooling attainment, with a focus on higher levels such as secondary and tertiary schools, where a gender gap is still likely to linger on in the present days. Besides family structures, among the correlates we explore are indicators for economic,
institutional, and religious factors. During this period, Italy indeed represents a microcosm which exhibits rich variations along all these dimensions. As a result, we present evidence that family structure is a robust correlate of the education gender gap during the historical period under examination, with a higher female to male ratio in enrollment in primary school being associated with nuclear residential habits and with equal partition of inheritance. Moreover, we find that the impact of family structure is persistent but, over contemporaneous data, only the latter dimension still matters. However, despite the fact that our results are consistent across different specifications, we should still be cautious in interpreting our results as causal.

The rest of the paper is organized as follows. In Section 2 we summarize the relevant literature. In Section 3 we present our newly-collected dataset. Section 4 contains the empirical analysis on the historical education gender gap. In Section 5 we turn to the contemporaneous education gender gap. We conclude in Section 6.

2. Literature

The literature on girls’ education in historical perspective has looked at multiple angles across time and space. Goldin (1998) and Goldin et al. (2006) study the evolution of secondary and tertiary education, respectively, for American women during the twentieth century.1 Becker and Wößmann (2008) investigate the determinants of girls’ education over nineteenth-century data for Prussia, with a focus on the positive influence of Protestantism.2 The education gender gap in Italy during the nineteenth-century is studied by Bertocchi and Bozzano (2013), who search for their deep, long-term determinants and find a beneficial influence on women’s outcomes of the medieval pattern of commerce, along the routes that connected Italian cities among themselves and with the rest of the world.3 In a more general setting, the long-term determinants of the gender gap have been studied by Galor and Weil (1996), who model the link between women’s inferior physical strength and their potential accumulation of human capital, and Alesina et al. (2013), who test the hypothesis that cultivation techniques requiring the plough determined the historical gender division of labor.4 With respect to this literature, our contribution is to focus on the influence of family culture on the education gender gap.

We shall present below four streams of the literature on gender roles, focusing respectively on family, development, history, and religion as potential determinants.

The potential role for family structure in determining women’s position has attracted a growing literature, following seminal work by Banfield (1958), Coleman (1990), and Todd (1990) on the general relevance of family ties for socio-economic outcomes. In particular, Todd develops a classification of family types organized along two leading axes: residential habits after marriage and inheritance rules. The former involve a distinction between “liberal” families, i.e., nuclear families

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1 Fernandez and Wong (2011) analyze the recent disappearance of the education gender gap in the US.
2 Protestantism was totally absent in Italy throughout the period under examination.
3 For Italy, see also Bertocchi et al. (2012), who document a reversal of the education gender gap in the post World War II period, and Bozzano (2014), who adapts the Global Gender Gap Index them to contemporaneous Italian data.
4 The plough system applied to the entire territory of Italy.
where children leave after marriage, and “authoritarian” families, i.e., extended (stem, or communitarian) families where different generations live together. In turn inheritance rules can be differentiated into “equal” and “unequal” ways to distribute family’s wealth, where inequality may involve primogeniture, unigeniture, or the absence of formal rules. The combination of these two axes leads to the following four family types: the absolute nuclear family, “liberal” and “unequal” (since no formal inheritance rules apply); the egalitarian nuclear family, “liberal” and “equal”; the stem family, “authoritarian” and “unequal”; and the communitarian family, “authoritarian” and “equal”. Todd also acknowledges the presence of a fifth type, the incomplete stem family, particularly widespread at the border between different national entities, which combines the “authoritarian” model with inheritance customs which are de jure but not de facto egalitarian. To determine the distribution of these family types within Western Europe Todd applies the “regressive method” due to Bloch (1949), which consists in starting with the most recent and complete data and from those to go back to historical and more fragmented data. In fact, it is from the patterns documented for the 1950s and 1960s on the basis of censuses that he can establish strong persistence of the observed family types back to the Middle Age.

Todd’s classification had been applied to the analysis of a number of issues. More relevant to this paper are those related to women’s position. Duranton et al. (2009) find that the European regions where egalitarian nuclear families tended to predominate have lower female participation in the labor force. Dilli et al. (2013) extend Todd’s classification to account for non-European parts of the world and find that it affects a gender equality index. An alternative measure of family structure is introduced by Alesina and Giuliano (2010, 2013), who construct a subjective variable measuring the strength of family ties using survey data and show that strong family ties are positively correlated with lower female labor force participation and more traditional attitudes towards women. Galasso and Profeta (2010) compare Todd’s classification with the one proposed by Alesina and Giuliano (2010, 2013) and find that they are broadly consistent.

A parallel study by Hajnal (1965, 1982) has focused on the European Marriage Pattern (EMP), a family model diffused in Northern Europe at least since the sixteenth century which involves nuclear residential patterns associated with late marriage. According to Hajnal these two features are closely connected: marriage can occur at an early age only in the presence of complex rather than nuclear families, since economic constraints would not allow the formation of an independent family at an early age. The pattern prevailing in Southern Europe, again according to Hajnal, was instead characterized by early marriage and complex families. The EMP has been associated with women’s position by De Moor and Van Zanden (2010), who argue that the practice of consensual marriage implied by this model led to distinctive cultural values based on equality between the sexes. This in turn encouraged women to work outside the household, contributing to economic growth. Along similar lines, Foreman-Peck (2011) emphasizes the implications of the higher marriage age for human capital investment for both sexes, but particularly for women, while

Voigtlander and Voth (2013) model and test the hypothesis that the Black Death triggered the shift toward the EMP by improving female employment prospects in the pastoral sector.

However, subsequent research has demonstrated that the nexus suggested by Hajnal between nuclear living arrangements and late marriage is not robust. Indeed, Southern Europe witnessed both early marriage, particularly for women, within nuclear families as well as late marriage within complex ones. Barbagli (1987) and Del Panta et al. (1996) show that examples of the former and the latter combination were diffused in Southern and Central Italy, respectively.7

This paper contributes to the literature on family structure and women’s outcomes by combining Todd’s and Hajnal’s classification and by highlighting the separate roles of residential habits, inheritance rules, and age at marriage.

Turning to the other potential correlates of the gender gap, the literature recently surveyed by Duflo (2012) looks at its nexus with development. With reference to the education dimension the author argues that, even though the gender gap in primary and secondary enrollment has rapidly gone down during the past decades, it still remains wider in poor countries. This evidence suggests that economic growth can have an important positive impact on gender equality but at the same time that it is generally not sufficient to overcome discrimination. The analysis in this paper confirms that economic factors indeed have a limited impact on gender equality in education, which justifies our focus on alternative explanations.

The influence of the pre-Unification political regimes on educational policies, including their gendered dimension, has been recognized by the history literature. Female education was largely neglected (Serristori, 1842) in most areas, with the exception of Lombardy, Rome, and the Duchy of Parma and Piacenza (Vigo, 1971). In many cases not only the quantity, but the quality of education for girls was inferior, being concerned primarily with knitting and sewing (Incatasciato, 1978). In the Bourbons’ territories the law explicitly allowed the hiring of illiterate female teachers (De Mauro, 1963). Unification brought the introduction of compulsory primary education for boys and girls and activated a regional convergence process in educational attainment, as well as the gradual reduction of the gender gap which is the focus of the present paper.8

Finally, the influence of religion on gender roles has been explored by Algan and Cahuc (2006) who, using survey data, show that Catholics, Orthodoxs, and Muslims are more prone to embrace the traditional male breadwinner conception than Protestants and atheists. Within the sociology literature, Esping-Andersen (1990) also associates a conservative view of women and the family with Catholic countries, while Bertocchi (2011) shows that Catholicism is negatively associated with the introduction of women suffrage in the 1870-1930 period. In this paper we aim at disentangling a potential role of religion even within a homogeneously Catholic country.

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7 Dennison and Ogilvie (2013)’s meta-study confirms that the European Marriage Pattern did not prevail throughout Europe and that its key components did not always coincide; it also casts doubts on the association between this family model and women’s empowerment, human capital accumulation, and economic success.

8 For further details see Vigo, 1971; Bertocchi and Spagat, 1997; Zamagni, 2002; A’Hearn et al., 2011; Bertola and Sestito, 2013.
3. Data and descriptive statistics

In order to study the long-term determinants of the education gender gap, we compile a database on education outcomes in late nineteenth-century Italy drawing from various sources. The main source is represented by data collected by the Italian Ministry of Agriculture, Industry and Commerce (MAIC). Our data cover 69 Italian provinces (at 1871 boundaries) in three points in time (1861, 1881, and 1901), spanning the initial forty years of the unified Kingdom of Italy. It should be stressed that enrollment data at the provincial level are only available until 1901. Data collection at this level of disaggregation is resumed only after World War II. For the same geographical entities, we complete the dataset with contemporaneous (1971 and 2001) census data on education outcomes by gender, taken from ISTAT (2009). We also collect data on the potential determinants of the historical and contemporaneous education gender gap. The variables are described below.

Education

For our dependent variables, we collect historical enrollment rates by gender for total population of primary school age,\(^9\) which consists of the population between 6 and 12 (excluded) years of age. On this basis we calculate a measure of gender equality given by the female to male ratio in enrollment rates in primary school.\(^10\) We also compute a separate measure at the lower and upper primary school level, only the former being compulsory.\(^11\) Beside data by gender, we also include general indicators of schooling defined as the number of pupils enrolled in primary school over the number of children of the corresponding 6-12 school age. Again we disaggregate this measure into a lower and an upper primary school level.

Our contemporaneous measures of education gender equality focuses on the secondary and tertiary levels, since equality is currently a well-established fact at the primary level. Rather than to enrollment, these data refer to the number of females and males holding high-school and college degrees, respectively.

Family structure

Our primary indicators for family structure reflect, separately, each of the two axes defined by Todd, i.e., residential habits and rules of inheritance. Accordingly, we first define a dummy variable that takes value 1 if the nuclear family pattern was prevailing in the province and 0 otherwise. Similarly, we define a dummy variable that takes value 1 if partible inheritance was prevailing and 0 otherwise. We follow Duranton et al. (2009) for the matching between the information in Todd and geographic units.

\(^9\) The lower level consisted of two years of schooling, extended to three after 1877; the upper level consisted of two years.

\(^10\) This measure is equivalent to the UN definition of the Gender Parity Index, where 0 corresponds to extreme inequality and 1 to full equality.

\(^11\) Because of lack of data at the provincial level, the denominators of the enrollment rates at the upper and lower levels are always the corresponding 6-12 population figures.
To account for Hajnal’s focus on age at marriage, we also collect information on female age at marriage, which is however available only at the regional level. Following Barbagli (1987) we take the percentage of women married below the age of 15 in 1872-75 (the original source is DIRSTAT, 1877).

Alternatively, we also account from Todd’s four-way classification (again matched to provincial units following Duranton et al., 2009). However, only three of the Todd family types are relevant for Italy, since the absolute nuclear family is never represented. Moreover, the stem family only appears in its incomplete form, in the area bordering the Austrian Empire. Furthermore, following Hajnal (1982), Barbagli (1987) and Del Panta et al. (1996), we augment Todd’s classification to distinguish for Italy between two different models of egalitarian nuclear family, which are respectively associated with late and early marriage age for women. As a result, we obtain for Italy four family types, generating a set of four dummies: egalitarian nuclear with late marriage, incomplete stem, communitarian, and egalitarian nuclear with early marriage.

*Economic development*

Our historical correlates for economic development involve measures of the level of wealth and the degree of structural transformation, both with a panel dimension over the 1861-1901 period. We proxy the former with the average height of conscripts aged 20. Data on average height for the three waves are taken from A’Hearn et al. (2009). Height is often employed in the economic history literature because it provides researchers with a measure of the stock of nutritional investment and therefore important indirect information on changes in the well-being of populations (Fogel et al., 1982; A’Hearn et al., 2011). To proxy for structural transformation we employ an index of industrialization computed by Ciccarelli and Fenoaltea (2013) and defined as the share of value added in manufacturing (excluding construction) over the share of the male population over age 15 (data are from census).


*Political regimes*

The period immediately preceding Unification, often referred to as Risorgimento, runs from the end of the Napoleonic era in 1815 until 1871 (when Rome becomes the capital). Before this historical turning point Italy was divided among four main dynasties: the Savoys, ruling roughly in the North West, i.e., in Piedmont, Liguria, and Sardinia; the Habsburgs, holding direct control over the Lombardy-Venetia, mostly located in today’s North East, as part of the Austrian Empire, as well as indirect control over the Granduchy of Tuscany and other minor Duchies nearby; the Pope,

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12 Since the available data neglect female stature, whose trend does not necessarily mimic that of males, as an alternative proxy of wealth we also consider population density data (from the Census) and estimates of per capita value added due to Istituto Guglielmo Tagliacarne (2011).
reigning over Rome and the Center South; and the Bourbons, in charge of the South and the Duchy of Lucca. To control for the legacy of deeply-rooted political institutions that may have shaped educational policies and the evolution and persistence of gender disparities, we construct a set of four dummies for the prevailing political regimes the 1815-1861 period, i.e., Savoys, Habsburgs, Pope, and Bourbons.

For the analysis of contemporaneous data we refer to four current macro-regions (North West, North East, Center, and South), which partially overlap with the historical political regimes.

**Religion**

While Italy has been historically a homogeneously Catholic country, we still aim at exploring the potential role of religion through various measures of the presence of the Church. Our first measure is the distance from Rome of the main city or town of each province. Second, we collect information from Palmieri (1857) about the number of parish churches in each province in 1857 and construct their density by dividing it by the provincial surface. Third, we take the fraction of secular teachers over the total number of teachers in 1862 (from MAIC, 1865 and 1874). Fourth, we measure the presence of the clergy using the number of priests (taken from Amati, 1860, and Correnti, 1858) over population in 1857, even though we have this information only at the regional level.

**Geography**

To capture residual factors such as the presence of different factor endowment, we also collect data on latitude, longitude, and the provincial fraction of plain terrain.

More details on the definitions and sources of all variables, as well as on methodological issues, can be found in the Appendix.

**Stylized facts**

Our dataset generates a set of stylized facts that can suggest preliminary hypotheses and considerations. The descriptive statistics reported in Table A2 reveal that in the 1861-1901 period the average enrollment rate in primary schools is 52.7 percent and the average female to male enrollment ratio is 0.83 (where 1 represents perfect equality). However there exists significant variation across provinces for both outcomes, ranging from 9.7 to 117.3 percent for the enrollment rate and from 0.05 to 1.51 for the female to male enrollment ratio. Moreover, the picture changes

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13 Becker and Wößmann (2009) employ distance from Wittenberg, Luther’s city, to measure the influence of Protestantism in late nineteenth century Prussia.  
14 Even though schooling was formally mandatory at the lower level, de facto the implementation of the law was gradual and patchy.  
15 The reason why enrollment rates can be greater than 100% is the presence of under-aged and/or over-aged enrollment. See the methodological notes in the Appendix for details.  
16 For some observations the ratio is larger than 1, which indicates a gender gap in favor of females.
considerably if one distinguishes between the lower and upper primary level: the former displays a female to male enrollment ratio of 0.85, for the latter the corresponding figure is 0.55. Figure 1 illustrates the time evolution of enrollment rates at the lower and upper level, showing a gradual increase over the years.

![Figure 1. Enrollment ratios at the lower and upper primary level, 1861-1901](image)

Authors’ elaboration from MAIC (various years).

![Figure 2. Female to male enrollment ratio at the lower and upper primary level, 1861-1901](image)
Figure 2 shows the female to male enrollment ratio, again for each level. For each year in the sample, gender disparity is more pronounced at the upper level. Over time, we observe a gradual closing of the gender gap. Nevertheless, for the last year of the sample differences between boys and girls are far from eradicated and remain especially deep at the upper level. In the following discussion, we focus on the female to male enrollment ratio at the upper primary level, where gender differences are more pronounced and more persistent.

The maps in Figure 3 again show the female to male ratio of the enrollment rates at the upper primary level, organized by quartiles, by province and over time. The darker is the colour of the provincial surface, the larger the female to male ratio. The maps confirm both the time trend and the differentiation across provinces.

![Figure 3. Female to male ratio upper primary enrollment ratio, 1861-1901](image-url)
The darker the colour, the higher is the female to male ratio upper enrollment rate in the province. Provinces are identified by abbreviations of current provincial names. See the Appendix for sources.

Figure 4 takes a glance at contemporary data. As of 1971, the position of females both at the secondary and the tertiary level is far from equal from that of males, especially at the tertiary level. The picture changes sensibly by 2001, with a near-closing of the gap at both levels, and in fact almost perfect equality at the tertiary level.

Figure 4. Female to male attainment ratio at the secondary and tertiary level, 1971 and 2001
Going back to the historical data, in the following discussion we relate our measure of gender equality in education to the potential determinants previously described, through a preliminary bivariate visual analysis.

We start by illustrating our data on family structure. Table A2 reveals that the nuclear family model prevails in about 53 percent of the provinces, while partible inheritance does so in about 87 percent. On average women marry before age 15 in only 3 percent of the regions, even though the figure is as high as 13 percent in the case of Sicily. The geographical distribution of the provinces where nuclear families and equal partition prevail is illustrated in Figures 5 and 6 respectively, while Figure 7 illustrates the geographic distribution of the four family types based on our adaptation of Todd’s classification. To be noticed is that equal partition and the incomplete stem family prevail in an area which is close, but not coincident, with the current North East.

Figure 5. Nuclear vs. extended family structure
Provinces are identified by abbreviations of current provincial names. See the Appendix for sources.

Figure 6. Equal vs. unequal partition of inheritance

Provinces are identified by abbreviations of current provincial names. See the Appendix for sources.

Figure 7. Family types
Provinces are identified by abbreviations of current provincial names. See the Appendix for sources.

Figure 8. Female to male enrollment ratio at the upper primary level, by residential habits, 1861-1901

Authors’ elaboration from MAIC (various years) and Todd (1990).

Figure 9. Female to male enrollment ratio at the upper primary level,
by inheritance rules, 1861-1901

In Figures 8 and 9 we report the female to male enrollment ratio at the upper primary level splitting the sample according to residential habits and inheritance rules, respectively. In this bivariate perspective women’s position is generally better when the nuclear family model and partible inheritance prevail.

Finally, Figure 10 reports our measure of gender equality organized according to Todd’s classification, as adapted to Italy. While the egalitarian nuclear family with late marriage seems to dominate in terms of women’s relative condition, a direct comparison across the other three models leads to less clear-cut conclusions.

Figure 10. Female to male enrollment ratio at the upper primary level,
Turning to our measures of development, Table A2 shows that the national average height, our proxy for prosperity, is about 163 centimeters. In Figure 11 we present the geographic distribution by quintiles of height in the 1861-1901 period. Again, the darker is the colour of the province, the higher is average provincial height. Over the forty years, height tends to increase but at different speeds, leaving large differences at the end of the period.
The darker the colour the higher is average height of conscripts in each year. Provinces are identified by abbreviations of current provincial names. See the Appendix for sources.

Next, Figure 12 illustrates the political geography of Italy according to our classification, while Figure 13 presents the time evolution of the female to male enrollment ratio in each of the pre-Unification geopolitical entities. In the provinces previously belonging to the Kingdom of Sardinia and in Lombardy-Venetia we observe larger female to male ratios. At the opposite we find the Bourbons’ territories, largely located in the South. These discrepancies can be explained, at least in part, by educational policies. For instance, schooling was compulsory until age 12 in Lombardy-Venetia. The political Unification of the country implied the adoption and implementation of national educational policies. At the very beginning of the Unification process, the 1859 Casati Law disciplined the school system by introducing the initial compulsory two-year level (extended to three years in 1877 by the Coppino Law) followed by a subsequent two-year level. The presence of free schools for boys and girls was made mandatory for the local authorities, in the effort to erase the regional differences inherited from the pre-Unitary polities and the differential supply and quality of primary schools. However, these legal provisions were enforced only slowly and
unevenly, so that spatial disparities are still present throughout the historical period under examination even though ultimately, by the end of the nineteenth century, these policies proved effective and contributed to a substantial decline, together with a regional convergence, of the gender gap in education.

Figure 12. Pre-Unification political regimes

Provinces are identified by abbreviations of current provincial names. See the Appendix for sources.

Figure 13. Female to male enrollment ratio at the upper primary level,
by pre-Unification political regimes, 1861-1901

In Figures 14 and 15 we illustrate two of the variables reflecting the influence of religion. Table A2 shows that the density of parish churches is 13 per squared kilometre on average, over a range of 1 to 70. It is highest in Sicily and some other Southern provinces. The proportion of secular teachers is on average almost 64 percent of total teachers, ranging from 15 to 98 percent, and is larger in the North and in Sardinia.
The darker the colour, the higher is the density of Parish churches present on the provincial territory. Provinces are identified by abbreviations of current provincial names. See the Appendix for sources.

Figure 14. Secular over total teachers

The darker the colour, the higher is the percentage of secular teachers over the total number of teachers in the province. Provinces are identified by abbreviations of current provincial names. See the Appendix for sources.

4. The historical education gender gap
In Table 1 we start with a set of OLS regressions (with clusters at the provincial level) where the dependent variable is the female to male ratio in enrollment rates\textsuperscript{17} at the upper primary level in the 1861-1901 period. To be recalled is that our dependent variable should be interpreted as an index of gender equality in education, so that a positive coefficient indicates that the corresponding regressor is associated with more gender equality and thus a smaller gender gap. In the first column we assess the effects of economic correlates by entering as regressors, beside time dummies, the 1861-1901 levels of height and industrialization, as proxies of prosperity and structural transformation: their coefficients are both significant and positive, which suggests that economic progress and modernization may indeed be associated with a more favourable position for women in the educational sphere. The (unreported) time dummies also display highly significant and positive coefficients.

In the next columns we add different sets of regressors, capturing political, religious, and cultural factors. Initially we add one set at a time. In Column 2 we introduce our main indicators for family structure, i.e., the dummy for nuclear families and the dummy for partition of inheritance. The former is associated with a more liberal family structure, as opposed to the authoritarian model, which in turn should favour a more advantageous position of women within the family and in society. The latter is associated with more equality among children including females, if compared to alternative customs such as male primogeniture, or discretion, which typically favour male heirs. As expected, both indicators are positive and significant.

In Column 3 we insert the set of dummies accounting for the political regimes prevailing before Unification. The reference is the Savoy dynasty which until 1861 ruled, roughly, the North West but then took charge of the unified kingdom of Italy. All the coefficients are negative, suggesting worse performances for the other regimes, even though only the coefficient for the Bourbons is significant, which signals that a deep divide exists from the start between the South and the rest of the country. To be noticed is that controlling for the pre-Unification political regimes means introducing regional fixed effect which non surprisingly reduce considerably the significance of the economic correlates.

In Columns 4-7 we control for different indicators for religion. The impact of the influence of the Catholic Church is a priori ambiguous. On the one hand, Catholicism is associated with a more traditional role of women, which implies a lesser need for them to work and to acquire more human capital through education. On the other, in the period under examination, education for boys and girls was frequently provided by religious institutions. We capture these factors with four variables: the density of parish churches, distance from Rome, the number of the members of the clergy over population, and the proportion of secular teachers. However, when entered one by one none of them is significant.\textsuperscript{18}

\textsuperscript{17} Enrollment rates are measured as the number of pupils of each gender over the school age (i.e., 6-12 years of age) population of the corresponding gender.

\textsuperscript{18} Same results obtain if they are entered together (we omit this specification).
The significance of nuclear family is even larger in a final regression in Column 8 where we enter together all the regressors which are significant in the previous specifications, i.e., we also control for political regimes beside economic correlates and time dummies. In this specification economic correlates are not significant, the inferior performance of the Bourbons is confirmed, and a superior performance emerges for the Habsburgs. The effect of family structure is sizeable, since in this last regression provinces with a predominance of nuclear families display an over 14% percent higher enrollment ratio. In other words, a one-standard-deviation increase of the presence of nuclear families is associated with an increase of 7 percentage points in the female to male ratio of the enrollment rate, which is equal to 12.9 percentage points in its sample averages. Moreover, the partial R-squared associated with nuclear families suggests that residually they explain 5.6 percent of the variation in the residual female to male enrollment ratio, conditional on the other controls. Similarly, the presence of partible inheritance implies an almost 17% percent higher enrollment ratio, while a one-standard-deviation increase of its presence is associated with an increase of 5.6 percentage points of the dependent (10.3 percentage points in its sample averages) and a partial R-squared of 6.1 percent.\(^{19}\)

| Table 1. The female to male enrollment ratio, Italy, 1861-1901: main determinants |
|------------------------------------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|-----------------|
| Dependent variable: Female to male ratio upper primary enrollment ratio |
| Height | 0.0275*** | 0.0582*** | -0.00632 | 0.0252*** | 0.0273*** | 0.0281*** | 0.0298*** |
|        | (0.00879) | (0.0104) | (0.0161) | (0.00922) | (0.00865) | (0.00894) | (0.0111) | (0.0141) |
| Industrialization | 0.121** | 0.0352 | 0.104* | 0.130** | 0.128** | 0.125** | 0.124** | -0.0150 |
|        | (0.0550) | (0.0522) | (0.0529) | (0.0583) | (0.0572) | (0.0557) | (0.0544) | (0.0500) |
| Nuclear family | 0.0764* | (0.0446) | 0.142*** |
| Partible inheritance | 0.177*** | (0.0499) | 0.168*** |
| Absburgs | -0.0121 | (0.0484) | 0.102* |
| Pope | -0.0668 | (0.0554) | 0.0200 |
| Bourbons | -0.220*** | (0.0691) | -0.180*** |
| Parish churches | -0.00178 | (0.00129) |
| Distance to Rome | -5.16e-05 | (0.000158) |
| Priests | 0.0226 | (0.0415) | -0.000646 |
| Secular teachers (%) | 0.48 | 0.53 | 0.53 | 0.48 | 0.48 | 0.48 | 0.48 | 0.60 |
| Observations | 207 | 207 | 207 | 207 | 207 | 207 | 207 | 207 |
| R-squared | 0.47 | 0.51 | 0.52 | 0.47 | 0.46 | 0.46 | 0.46 | 0.58 |

Robust standard errors in parentheses clustered at the provincial level. *** p<0.01, ** p<0.05, * p<0.1. A constant and time fixed effects are always included. The omitted dummy for political regimes in Column 2 is for the Savoys.

\(^{19}\) In a further specification we also control for geography (latitude, longitude, and plain terrain fraction) which should capture factor endowments, but all dimensions are insignificant and do not affect our results.
Similar results (not reported for brevity) obtain if height is replaced by alternative proxies of prosperity, i.e., the 1861-1901 levels of population density and estimates of per capita value added due to Istituto Guglielmo Tagliacarne (2011). We also obtain very similar results if we replace the dependent variable with the aggregate primary level and at the lower level. However, the effect of the political regimes becomes insignificant and the effect of the dummies for family structure is slightly weaker, which is not surprising since the upper level is non-compulsory and therefore more likely to reflect differences reflecting individual choices, while at the lower level and at the aggregate we should expect less variability in all the relevant dimensions. Moreover, a trace of the effect of religion emerges at the lower level, where the presence of the clergy becomes beneficial for relative female enrollment, even though this result should be taken with caution since this variable is measured at the regional rather than provincial level.

In Table 2 we explore the impact of family structure in further detail. In Column 1 we add to the basic regression with economic correlates and time dummies the proportion of women married before age 15. In Column 2 we also add the previous two indicators of family structure. Age at marriage is an important indicator of the degree of subordination of women with respect to men and therefore of gender inequality. As expected, early marriage exerts a negative influence on relative female enrollment in both specifications, while the indicator for nuclear family is finer larger and more precisely measured. In Column 3 we insert three dummies reflecting Todd’s classification (1990), augmented to account for differential female age at marriage within nuclear families. The omitted dummy is the one for the egalitarian nuclear family with late female age at marriage. The coefficients for the egalitarian nuclear family with early marriage, the communitarian family, and the incomplete stem family are all significantly negative. However, when in Column 4 we also add political regimes to reproduce the extended specification in Table 1, Column 8, the additional explanatory power of marriage age disappears, while the other two dimensions capturing residential and inheritance patterns are confirmed as significant (to be recalled however if that the variable reflecting age is measured at the regional rather than provincial level). Similarly, in the extended specification in Column 5 only the incomplete stem family – which involves complex residential habits as well as de facto primogeniture – retains an explanatory power, possibly because of the partial overlap between political regimes and prevailing family structures. This exercise confirms that the two separate dimensions we choose to emphasize from the start not only display a more robust influence but also allow a more precise assessment of which is the channel through which family structure exerts its influence.
Table 2. The female to male enrollment ratio, Italy, 1861-1901: further indicators for family structure

<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
<th>(2)</th>
<th>(3)</th>
<th>(4)</th>
<th>(5)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Dependent variable: Female to male ratio upper primary enrollment ratio</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Height</td>
<td>0.0123</td>
<td>0.0432***</td>
<td>0.0432***</td>
<td>0.0221</td>
<td>0.0205</td>
</tr>
<tr>
<td></td>
<td>(0.0127)</td>
<td>(0.0127)</td>
<td>(0.0137)</td>
<td>(0.0143)</td>
<td>(0.0143)</td>
</tr>
<tr>
<td>Industrialization</td>
<td>0.138***</td>
<td>0.0375</td>
<td>0.0142</td>
<td>-0.0159</td>
<td>-0.00885</td>
</tr>
<tr>
<td></td>
<td>(0.0545)</td>
<td>(0.0503)</td>
<td>(0.0471)</td>
<td>(0.0512)</td>
<td>(0.0501)</td>
</tr>
<tr>
<td>Early marriage</td>
<td>-0.0109**</td>
<td>-0.0140***</td>
<td>0.00101</td>
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</tr>
<tr>
<td></td>
<td>(0.00525)</td>
<td>(0.00480)</td>
<td>(0.00528)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Nuclear family</td>
<td>0.110***</td>
<td>0.141***</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.0455)</td>
<td>(0.0527)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Partible inheritance</td>
<td>0.163***</td>
<td>0.167***</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.0488)</td>
<td>(0.0568)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Egalitarian nuclear fam. early marriage</td>
<td>-0.139***</td>
<td>0.121</td>
<td></td>
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<td></td>
</tr>
<tr>
<td></td>
<td>(0.0502)</td>
<td>(0.0853)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Incomplete stem family</td>
<td>-0.282***</td>
<td>-0.279***</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.0546)</td>
<td>(0.0666)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Communitarian family</td>
<td>-0.136***</td>
<td>-0.0932</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.0459)</td>
<td>(0.0727)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Absburgs</td>
<td></td>
<td>0.102*</td>
<td>0.0759</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td></td>
<td>(0.0572)</td>
<td>(0.0619)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Pope</td>
<td>0.0202</td>
<td></td>
<td>-0.0322</td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.0698)</td>
<td></td>
<td>(0.0850)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Bourbons</td>
<td>-0.186***</td>
<td>-0.296***</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td></td>
<td>(0.0649)</td>
<td>(0.100)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Observations</td>
<td>207</td>
<td>207</td>
<td>207</td>
<td>207</td>
<td>207</td>
</tr>
<tr>
<td>R-squared</td>
<td>0.49</td>
<td>0.55</td>
<td>0.55</td>
<td>0.60</td>
<td>0.60</td>
</tr>
<tr>
<td>Adj. R-squared</td>
<td>0.48</td>
<td>0.53</td>
<td>0.54</td>
<td>0.58</td>
<td>0.58</td>
</tr>
</tbody>
</table>

Robust standard errors in parentheses clustered at the provincial level. *** p<0.01, ** p<0.05, * p<0.1. A constant and time fixed effects are always included. The omitted dummies are the egalitarian nuclear family with late marriage for Todd’s family types in Column 3 and 5 and the Savoys for political regimes in Column 5.

In Table 3 we try to evaluate whether the influence of family structure varies throughout the period under consideration, by adding to the extended specification of Table 1, Column 8 a set of interactions between time and each of the two indicators for residential patterns and inheritance rules. The fact that the interactions are not significant and the positive coefficient is retained for both indicators confirms their lasting role throughout the historical period under investigation.

Table 3. The female to male enrollment ratio, Italy, 1861-1901: dynamics

<table>
<thead>
<tr>
<th></th>
<th>(1)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Dependent variable: Female to male ratio upper primary enrollment ratio</td>
<td></td>
</tr>
<tr>
<td>Height</td>
<td>0.0223*</td>
</tr>
</tbody>
</table>
We can therefore conclude that, over the first four decades after Unification, economic and religious factors are not exerting a significant influence on the education gender gap, geopolitical factors do so but only marginally, while nuclear residential patterns and equal inheritance rules are strongly and persistently associated with a lower gap. Even though we should be cautious in interpreting our results as causal, the evidence we present is robust across different specifications.

5. The contemporaneous education gender gap

In this section we investigate the long-term persistence of the effect of family structure on contemporaneous measures of the education gender gap. Since gender equality at the primary level is a well-established fact in post-WW2 Italy, we focus on higher educational levels, i.e., secondary and tertiary levels. As shown in Figure 4, the position of females at both levels, as of 1971, is far from equal from that of males, especially at the tertiary level. The picture changes sensibly in 2001, with a near-closing of the gap at both levels, and in fact almost perfect equality at the tertiary level, which suggests the interplay of different and level-specific factors in explaining the gap and its evolution. Within this contemporaneous dataset, family structure is measured again as within the previous historical dataset since, while it is true that it has evolved, we aim at capturing its long-term influence, while avoiding at the same time likely endogeneity problems which would be associated with updated contemporaneous measures. In Table 4 we present our results, for a
specification which includes, for each year, the corresponding income level, the variables reflecting family structure, and dummies for the macro-regions that reflect the current political geography of Italy (with the North West, roughly corresponding to the Savoys’ territories, as a reference). Column 1 shows that partible inheritance is positively associated with relative female attainment at the secondary level in 1971, while the presence of nuclear families no longer matters.\textsuperscript{20} The same conclusion can be reached for the tertiary level in the same year and again for the secondary and tertiary in 2001 (Columns 3, 5 and 7).

When we turn to Todd’s set of dummies we find that, relative to the egalitarian nuclear family with late marriage, the coefficient of the egalitarian nuclear with early marriage is significantly negative in three out of four regressions, while that of the communitarian family is positive for the secondary level in 1971. Taken together, these results are quite different from those emerging from the historical measures of the gender gap, and even harder to interpret.

To sum up, over the long run a single dimension of family structure, the one reflecting equality in inheritance rule, exerts a persistent influence on the gender gap in education, while residential patterns no longer matter.

\textsuperscript{20} In their analysis of contemporaneous welfare systems Galasso and Profeta (2013) also find an impact of partible inheritance but not of residential habits.
### Table: Family Structure across Italian Provinces

<table>
<thead>
<tr>
<th>Macro-region</th>
<th>Parameter 1</th>
<th>Parameter 2</th>
<th>Parameter 3</th>
<th>Parameter 4</th>
<th>Parameter 5</th>
<th>Parameter 6</th>
<th>Parameter 7</th>
<th>Parameter 8</th>
</tr>
</thead>
<tbody>
<tr>
<td>North East</td>
<td>0.0716**</td>
<td>0.0283</td>
<td>0.0993****</td>
<td>0.00751</td>
<td>0.0255</td>
<td>0.0774</td>
<td>0.00273</td>
<td>0.0386</td>
</tr>
<tr>
<td></td>
<td>(0.0354)</td>
<td>(0.0710)</td>
<td>(0.0327)</td>
<td>(0.0362)</td>
<td>(0.0216)</td>
<td>(0.0609)</td>
<td>(0.0243)</td>
<td>(0.0593)</td>
</tr>
<tr>
<td>Center</td>
<td>0.00531</td>
<td>-0.0106</td>
<td>0.0630****</td>
<td>-0.00716</td>
<td>0.0230</td>
<td>0.0805</td>
<td>0.0657***</td>
<td>0.0800</td>
</tr>
<tr>
<td></td>
<td>(0.0293)</td>
<td>(0.0564)</td>
<td>(0.0225)</td>
<td>(0.0317)</td>
<td>(0.0213)</td>
<td>(0.0561)</td>
<td>(0.0225)</td>
<td>(0.0583)</td>
</tr>
<tr>
<td>South</td>
<td>-0.0123</td>
<td>0.0360</td>
<td>0.0284</td>
<td>0.00802</td>
<td>0.0524</td>
<td>0.139**</td>
<td>0.114***</td>
<td>0.172**</td>
</tr>
<tr>
<td></td>
<td>(0.0444)</td>
<td>(0.0559)</td>
<td>(0.0277)</td>
<td>(0.0303)</td>
<td>(0.0390)</td>
<td>(0.0624)</td>
<td>(0.0424)</td>
<td>(0.0663)</td>
</tr>
</tbody>
</table>

Observations: 69  
R-squared: 0.29  
Adj. R-squared: 0.22

Robust standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. A constant is always included. The omitted dummy for macro-regions is North West.

### 6. Conclusion

The first achievement of this paper is to adapt existing classifications of family structure to the specific historical circumstances prevailing in Italy at least since the nineteenth century. The second achievement of the paper is to show that family structure exhibits significant variation across Italian provinces, despite the fact that in cross-sectional studies the country is normally viewed as homogenous in this respect. Finally, the paper shows that family structure affects the gender gap in education in a lasting fashion, even after controlling for a rich set of correlates reflecting economic performances, the political geography of the country prior to Unification, and the influence of the Catholic Church.

Over a century and a half after the Unification of Italy in 1861, the education gender gap has been gradually closing, and even reversing in recent years, following a pattern which is common to most developed countries. As women’s lives have been transformed thanks to the improvement of their educational attainment together with their economic and social position, the structure of the family inherited from history has also been transformed, with a rapid decline of marriage and a drastic reduction of fertility. To understand how family structure and the gender gap are jointly evolving is the next challenge for future research.

### References


APPENDIX

This appendix describes the data underlying our study in more detail. The dataset contains information from 69 Italian provinces at 1871 boundaries. We collect province-level data at three points in time (1861, 1881, 1901) whenever available, spanning the first forty years after Italian Unification. Table A1 contains variable definitions and sources. Methodological notes on the construction of the dataset follow next. Table A2 presents summary statistics.

Table A1. Variable definitions and sources

<table>
<thead>
<tr>
<th>Variable</th>
<th>Construction</th>
<th>Main Source</th>
</tr>
</thead>
<tbody>
<tr>
<td>Primary enrollment rate</td>
<td>Pupils enrolled in primary school/population of primary school age (6-12) *100</td>
<td>MAIC (various years)</td>
</tr>
<tr>
<td>Lower primary enrollment rate</td>
<td>Pupils enrolled in lower primary school/ population of primary school age (6-12) *100</td>
<td>MAIC (various years)</td>
</tr>
<tr>
<td>Upper primary enrollment rate</td>
<td>Pupils enrolled in upper primary school/ population of primary school age (6-12) *100</td>
<td>MAIC (various years)</td>
</tr>
<tr>
<td>Male primary enrollment rate</td>
<td>Male pupils enrolled in primary school/ male population of primary school age (6-12) *100</td>
<td>MAIC (various years)</td>
</tr>
<tr>
<td>Female primary enrollment rate</td>
<td>Female pupils enrolled in primary school/ female population of primary school age (6-12) *100</td>
<td>MAIC (various years)</td>
</tr>
<tr>
<td>Variable</td>
<td>Formula</td>
<td>Source</td>
</tr>
<tr>
<td>----------</td>
<td>---------</td>
<td>--------</td>
</tr>
</tbody>
</table>
| Female to male primary enrollment ratio | \[
\frac{\text{female pupils enrolled in primary school}}{\text{female population of primary school age (6-12)}} / \frac{\text{male pupils enrolled in primary school}}{\text{male population of primary school age (6-12)}} = \frac{\text{female pupils enrolled in primary school}}{\text{male pupils enrolled in primary school}} \times \text{masculinity of primary school age (6-12) population}
\] | MAIC (various years) |
| Female to male upper primary enrollment ratio | Female pupils enrolled in upper primary school/ male pupils enrolled in upper primary school * masculinity of primary school age (6-12) population | MAIC (various years) |
| Female to male lower primary enrollment ratio | Female pupils enrolled in lower primary school/ male pupils enrolled in lower primary school * masculinity of primary school age (6-12) population | MAIC (various years) |
| Female to male secondary attainment ratio 1971 | Females with at least a high-school diploma/ males with at least a high-school diploma. Year 1971. | ISTAT (2009) |
| Female to male tertiary attainment ratio 1971 | Females with a university degree/ males with a university degree. Year 1971. | ISTAT (2009) |
| Female to male secondary attainment ratio 2001 | Females with at least a high-school diploma over female population over age 19/ males with at least a high-school diploma over male population over age 19. Year 2001. | ISTAT (2009) |
| Female to male tertiary attainment ratio 2001 | Females with a university degree over female population over age 19/ males with a university degree over male population over age 19. Year 2001. | ISTAT (2009) |
| Masculinity | Male population of primary school age (6-12)/ female population of primary school age (6-12) | Census (various years) |
| Latitude | Latitude of the main city of the province in decimal degrees. | Passim |
| Longitude | Longitude of the main city of the province in decimal degrees. | Passim |
| Plain terrain | Percentage of plain terrain over the total surface of the province | Passim |
| Distance from Rome | Distance of the main provincial city or town to Rome, in kilometres (great-circle distance). | http://distanzechilometriche.net |
| Clergy | Number of clergy members/ 100 individuals in population, in year 1857 (at the regional level) | Amati (1860), Correnti (1858) |
| Parish churches | Number of parish churches in existence in 1857/ provincial surface in squared kilometers | Palmieri (1857) |
| Secular teachers | Number of secular teachers/ total number of teachers, in year 1862 | MAIC (1865), MAIC (1874) |
| Nuclear family | Dummy variable: 1 if the nuclear family type prevails, 0 otherwise | Todd (1990) |
| Partible inheritance | Dummy variable: 1 if the partible inheritance rule prevails, 0 otherwise | Todd (1990) |
| Early marriage | Women below age 15/ married women, in 1872-75 (at the regional level) | Barbagli (1987), DIRSTAT (1877) |
| Egalitarian nuclear family with early marriage | Dummy variable: 1 if the egalitarian nuclear family type with early marriage prevails, 0 otherwise | Todd (1990), Del Panta et al. (1996) |
| Egalitarian nuclear family with late marriage | Dummy variable: 1 if the egalitarian nuclear family type with late marriage prevails, 0 otherwise | Todd (1990), Del Panta et al. (1996) |
| Incomplete stem family | Dummy variable: 1 if the incomplete stem family type prevails, 0 otherwise | Todd (1990), Del Panta et al. (1996) |
| Communitarian family | Dummy variable: 1 if the communitarian family type prevails, 0 otherwise | Todd (1990), Del Panta et al. (1996) |
| Savoys | Dummy variable: 1 if the province was ruled by the Savoy-Carignano dynasty prior to Unification, 0 otherwise | Pécout (1999) |
| Habsburgs | Dummy variable: 1 if the province was ruled by the Habsburg dynasty prior to Unification, 0 otherwise | Pécout (1999) |
| Pope | Dummy variable: 1 if the province was ruled by the Pope prior to Unification, 0 otherwise | Pécout (1999) |
| Bourbons | Dummy variable: 1 if the province was ruled by the Bourbon dynasty prior to Unification, 0 otherwise | Pécout (1999) |
| Height | Average height at age 20 of military conscripts | A’Hearn et al. (2011) |
| Industrialization | Provisonal index of relative industrialization calculated on the basis of census data as the share of industrial value added, excluding construction, over the share of the male population over age 15 | Ciccarelli and Fenoaltea (2013) |

Methodological notes

_Historical education data:_ We collect data for 1861, 1881, and 1901. Due to historical circumstances, we encountered a few problems because of lack of data in particular referring to the first wave. This arises from the fact that the Italian Unification process was a gradual process and therefore official statistical annuals could not take into account those provinces which entered the Italian Kingdom only in subsequent years. As a result data on gross enrollments (i.e., number of pupils of any age who are enrolled in primary schools as a percentage of the total official school-age children) are constructed from various sources and mainly from Statistica del Regno d'Italia, Istruzione pubblica e privata, a statistical series compiled and published by the Ministry of Agriculture, Industry and Commerce (MAIC) from 1865 to 1906, and from Vigo (1971). For most provinces we impute to 1861 enrollment data referring to 1862-63. For provinces belonging to today’s Veneto data on enrollments for the first wave refer to 1856 and are from Buonazia (1870) and Vigo (1971) for primary school and Sacchi (1858) for upper primary school. For Mantua data refer to 1851 and are from Vigo (1971) for primary school and from MAIC (1872) for upper primary school. For Rome (and Comarca) data refer to 1858 and are from Vigo (1971) for primary school and from MAIC (1872) for upper primary school.

To determine enrollment rates, we consider 6-12 (excluded) as the relevant primary school age. In calculating lower and upper primary enrollment rates we still employ the population 6-12
(excluded) because we have no data for the specific grade ages. Therefore at the lower and upper level we have an unadjusted measure of enrollment (Benavot and Riddle, 1988) which induces a downward bias in the measure of enrollment rates.

In some cases gross enrollment rates are greater than 100% and this is due to under-aged and/or over-aged enrollment with respect to the school age being considered (and this is even more compelling during a period of expansion of mass education with a high incidence of late entrance). However, information on net rates (defined as the ratio between enrollments of age 6-12 over the population of age 6-12) is not available.

**Family structure:** Our first indicator is a dummy variable taking value 1 if the province was dominated by nuclear residential habits, 0 if complex ones. Our second indicator is a dummy variable taking value 1 if the province was dominated by equal partition rules of inheritance, 0 if by unequal ones. Information on both indicators is from Todd (1990). Our third indicator is the percentage of women married below the age of 15 in 1872-75, taken from Barbagli (1987) (the original source is DIRSTAT, 1877). Finally, we adapt the classification proposed by Todd (1990) to the Italian specific regional differentiation following Hajnal (1982), Barbagli (1987), and Del Panta et al. (1996) and code provinces according to a set of four dummies, one for each family type: egalitarian nuclear with late marriage, incomplete stem, communitarian, and egalitarian nuclear family with early marriage.

**Contemporaneous education data:** We collect data on female and male secondary and tertiary attainment for 1971 and 2001 from national censuses (ISTAT, 2009) and we obtain female to male ratios to be employed as our contemporaneous dependent variables. For 1971 our dependent variable is not weighted by relative population whereas for 2001 our variables are weighted by female and male population aged 19 or more.

**Religion:** We measure the influence of religion through four variables. Our first measure is if the distance from Rome of the main city or town of each province in kilometers. The source is http://distanzechilometriche.net. Second, we collect information from Palmieri (1857) about the number of parish churches in 1857 and obtain their density by dividing it by the provincial surface in squared kilometers. Third, we take the fraction of secular teachers over the total number of teachers in 1862. The main source is MAIC (1865), and MAIC (1874) for provinces of Veneto, Mantua, and Rome. Fourth, we measure the number of clergy in 1857 (taken from Amati, 1860, and Correnti, 1857) per 100 individuals in the population. This fourth variable is only available at the regional, rather than provincial, level.

**Ruling dynasties:** This set of binary dummy variables are intended to capture the ruling dynasties immediately prior to Unification: the Savoy-Carignano, who ruled in the North Western areas of Italy, i.e., Piedmont, Liguria, and the island of Sardinia; the Habsburgs, who held direct control over Lombardy-Venetia, mostly in the North East, as part of the Austrian Empire, as well as indirect control over the Granduchy of Tuscany (Habsburg-Lorena) and other minor Duchies nearby (the Duchy of Modena under Este of Austria, the Duchy of Parma and Piacenza, and the Duchy of Massa Carrara under Habsburg-Lorena); the Pope, reigning over Rome and the Center South; and
the Bourbons, in charge of the South and the Duchy of Lucca. Even though the latter falls under the Habsburgs in 1847, we consider it as dominated by the Bourbons throughout the 1816-1861 period. Information is from Pécout (1999).

**Height:** Data about wealth of ancient Italian provinces are not available. For this reason we employ height as a proxy for wealth. In the economic history literature it is common to make use of anthropometric measures because they provide important indirect information on changes in the well-being of the population. Data on height refer to the mean height of military conscripts aged 20 at the provincial level and are taken from A’Hearn et al. (2011, Table S3). For an exhaustive presentation of the use of this kind of anthropometric data as a plausible proxy for well-being and living standards see A’Hearn et al. (2011, ch. 2). For a discussion about the methodological approach employed to estimate average height see A'Hearn et al. (2009).

**Industrialization:** Data on industrialization at the provincial level are taken from Ciccarelli and Fenolleta (2013, Table 2). The variable is a provincial index of relative industrialization calculated on the basis of census data as the share of industrial value added, excluding construction, over the share of the male population over age 15. Ciccarelli and Fenolleta (2013) explain that using this age/sex group provides the best proxy of local GDP, because adult males represented the most mobile and flexible segment of the labor force and also because male labor force participation rates are uniformly high and not clouded, as female rates are, by varying local conventions. Data for the first wave refer to 1871.

### Table A2. Summary statistics

<table>
<thead>
<tr>
<th>Variable</th>
<th>Obs</th>
<th>Mean</th>
<th>Std. Dev.</th>
<th>Min</th>
<th>Max</th>
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<tr>
<td>Enrollment rate</td>
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<td>Female enrollment rate</td>
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<td>Female to male ratio secondary attainment 1971</td>
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<td>Mean 2</td>
<td>Mean 3</td>
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