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# State Dependence in Welfare Receipt: Transitions Before and After a Reform

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# State Dependence in Welfare Receipt: Transitions Before and After a Reform

### **Abstract**

We study state dependence in welfare receipt and investigate whether welfare transitions changed after a welfare reform. Using data from the German Socio-Economic Panel, we apply dynamic multinomial logit estimators and find that state dependence in welfare receipt is not a central feature of the German welfare system. We find that welfare transitions changed after the reform: transitions from welfare to employment became more likely and persistence in welfare and inactivity declined. We observe a large relative increase in transitions from employment to welfare. Immigrants' responsiveness to the labor market situation increased after the reform.

JEL-Code: I380, J610.

Keywords: social assistance, state dependence, unemployment benefit II, immigration, dynamic multinomial logit.

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#### 1 Introduction

In recent years, the German labor market was surprisingly robust to the Great Recession of 2009.<sup>1</sup> Interestingly, this positive development was preceded by a substantial reform of the German welfare system that aimed at enhancing work incentives and labor market flexibility, particularly for the unemployed and for welfare recipients. Some economists assert that the welfare reform contributed to the positive development as it may have affected welfare transitions (e.g., SVR 2011). In particular, the reform provided incentives to exit welfare and to reduce welfare persistence.

This paper studies the structural determinants of welfare transitions. We are particularly interested in true state dependence as an indicator of the extent to which the experience of welfare receipt causally affects subsequent welfare receipt. We address the following research questions: (1) is there true state dependence in the German welfare system? Comparing the outcomes before and after the recent reform, we provide a comprehensive picture of welfare transitions for the entire population. We do not focus on causal reform effects which may only be identifiable for subgroups. Instead, we ask: (2) did the patterns and dynamics of welfare transitions change from before to after the reform, and (3) are welfare transitions more responsive to the labor market situation after the reform? In addition, we examine heterogeneities in welfare transitions and in changes in welfare transitions for immigrants and natives because the literature shows that the patterns of welfare receipt differ for these groups (Barrett and McCarthy 2008).

Our study of welfare transitions ties in with an international literature on state dependence in welfare receipt which typically applies dynamic discrete choice models. Closest to ours is the contribution by Hansen and Lofstrom (2009) who use data on Swedish men to study the transition between welfare receipt, unemployment, and employment. They find higher state dependence among immigrants than natives. Hansen and Lofstrom (2011) separately study welfare exit and

<sup>&</sup>lt;sup>1</sup> Between 2005 and 2011, the unemployment rate dropped from 13.0% to 7.9% and employment surged from 38.9 to 41.1 millions (SVR 2012).

entry of Swedish natives and immigrants and find that the difference in welfare receipt between natives and immigrants results from differences in entry to rather than in exit from welfare. In the literature on welfare state dependence, Hansen et al. (2006) analyze Canadian welfare participation and Chay et al. (2004) focus on Californian data. Both studies find true state dependence in welfare receipt which varies across population groups. This is confirmed for the case of the U.K. by Cappellari and Jenkins (2008), who find evidence for state dependence but point to its heterogeneity across subsamples.

Knowledge about state dependence in welfare receipt may contribute to the explanation of the German "job miracle". While Burda and Hunt (2011) see employer expectations, wage moderation, and working time accounts as the key factors behind the German job miracle, SVR (2011) also discuss the role of recent reforms. We provide evidence on the plausibility of the connection between the 2005 reform and subsequent employment dynamics.<sup>2</sup> Given that many countries are faced with high unemployment, the German experience may provide an informative benchmark case.

This study relates to the international literature on the consequences of welfare reforms for recipient behaviors. Following the 1996 reform of the U.S. federal welfare program, studies addressed a variety of outcomes (for a survey, see Blank 2002), among them the propensity to take up work (e.g., Grogger and Karoly 2005), responses to time limited eligibility, and the relevance of the macroeconomy for labor force participation (Bitler and Hoynes 2010, Ziliak et al. 2000). Blank (2002) summarizes evidence of substantial changes in welfare transition patterns

<sup>&</sup>lt;sup>2</sup> Some contributions already study the response of the German labor market to the recent welfare reforms. Fahr and Sunde (2009) and Klinger and Rothe (2012) find that the early Hartz Reforms significantly improved the efficiency of labor market matching, benefiting particularly the long term unemployed. Several studies evaluated the effects of elements of the reform packages (e.g., Huber et al. 2011). Caliendo and Hogenacker (2012) summarize that labor market institutions became more efficient and work incentives for the unemployed increased after the reform.

in response to the U.S. welfare reforms.<sup>3</sup> By comparing welfare dynamics before and after the German reform, we contribute to this literature.

In addition, we contribute to the literature on the connection between individual welfare transitions and aggregate labor market conditions (e.g., Hoynes 2000 and Hoynes et al. 2012). Hoynes (2000) showed the close connection between unemployment and welfare receipt in the United States between 1987 and 1992 when welfare recipients strongly responded to job opportunities and wage growth.

Using data from the German Socio-Economic Panel Study (SOEP), we estimate dynamic multinomial logit models to study transitions in the German welfare system. Our results show little evidence of true state dependence. We find that transition patterns changed after the reform: transitions to employment became more likely and persistence in welfare and inactivity declined. We observe a large relative increase in transitions from employment to welfare. Among immigrants, welfare persistence and welfare-to-employment transitions became more responsive to the labor market after the reform. Compared to standard bivariate estimation approaches which dominate the literature on welfare transitions (e.g., Hansen et al. 2006 or Cappellari and Jenkins 2009), our multinomial outcome measure allows us to differentiate alternative pathways in and out of welfare benefit receipt.

This paper is structured as follows. In section two we summarize the institutional framework and the key reform elements that might affect state dependence. Section three describes the data and section four the empirical approach. The results are shown in section five and section six presents concluding remarks.

<sup>&</sup>lt;sup>3</sup> For evidence on reforms in the U.K., Sweden, and Canada see, e.g., Brewer et al. (2006), Edmark (2009), and Fortin et al. (2004).

#### 2 Institutions

In response to increasing unemployment rates the German government implemented far-reaching reforms which changed the welfare and the unemployment insurance system (Schneider 2012). Because it cut back on some claims against the welfare state the reform received substantial public attention and opposition. Next, we describe the German welfare system before and after the reform and discuss why welfare transitions might have changed. Then, we briefly describe the situation of immigrants.

In case of unemployment, workers are generally covered by the unemployment insurance. Unemployment benefits (*Arbeitslosengeld*) replace up to 67% of previous net earnings. The reform reduced the maximum duration of benefit payment from 32 to 24 months. The benefit is now labeled unemployment benefit I (UB I). Prior to the reform, those who had exhausted their unemployment benefit entitlement and those who were not (yet) entitled to unemployment benefits were eligible for unemployment assistance (*Arbeitslosenhilfe*), a tax-financed means-tested transfer. Unemployment assistance replaced up to 57% of prior net earnings. Also, prior to the reform, households could claim social assistance (*Sozialhilfe*) if their total income—independent of its source—fell below the legally defined subsistence level. Social assistance was a means-tested program that, in case of need, could also be paid in addition to labor earnings and unemployment benefits.

The reform then combined unemployment assistance and social assistance in the so-called unemployment benefit II (UB II), a means-tested and tax-financed benefit. Households in need of support may be eligible for UB II, independent of whether their members are employed, receive UB I, or exhausted their UB I eligibility. The benefit covers the legally defined minimum income and is not related to prior earnings. Households in need can claim UB II if their adult members are able to work at least 15 hours per week. If the household members are not able to work,

e.g., due to sickness or disability, the household is—as before—entitled to social assistance. The means test is administered at the household level.

The reform came into effect in January 2005 as one element of a wider reform project. The reform project had several objectives: (a) to improve the effectiveness and efficiency of labor market services. Thus, after the reform, local employment offices introduced differentiated approaches to support the unemployed at an individual level. (b) To activate the unemployed based on the idea of 'fordern and fördern', i.e., 'to assist and demand.' Since the reform, the employment offices explicitly demand individual effort and have the unemployed sign 'agreements on objectives.' At the same time search incentives were increased by shorter durations of unemployment benefit payouts and by an intensified use of sanctions. (c) Finally, labor market regulations were relaxed, e.g., with respect to employment protection, temporary employment, and temporary agency employment (Klinger and Rothe 2012, Caliendo and Hogenacker 2012).

The reform adjusted the regulations of earnings allowances and marginal tax rates to increase work incentives (see, e.g., Dietz et al. 2011): the maximum earnings allowance increased and marginal tax rates declined.<sup>4</sup> In addition to strengthening work incentives, the reform also requires welfare recipients to actively search for jobs: all recipients of UB II have to look for a job and are obliged to discuss their search strategy with the employment office. In contrast, before the reform social assistance benefits were paid independent of labor market status and search effort. These changes may well reduce welfare persistence and state dependence in welfare.

Immigrants are treated like natives within the unemployment insurance, i.e., with respect to unemployment benefit and unemployment assistance before the reform and UB I after the reform. Their eligibility depends on the individual contribution record. The situation for immigrants is more complex in the minimum income support programs of social assistance and UB II. Individ-

<sup>&</sup>lt;sup>4</sup> For details see Riphahn and Wunder (2013) where we compare the characteristics associated with benefit receipt among natives and immigrants and provide a non-parametric study of the groups' respective life cycle trajectories of benefit receipt. The paper does not look at state dependence and does not provide a dynamic analysis of the situation before and after the reform.

uals without German citizenship can receive minimum income support if they are (i) permanently in Germany, (ii) physically able to work (after the reform), and (iii) potentially allowed to take up employment; the last condition excludes, e.g., asylum seekers. Ethnic Germans (*Aussiedler*) as well as naturalized immigrants are treated like natives.<sup>5</sup> Immigrants residing in Germany in order to find employment are generally not eligible for benefits. However, a long list of circumstances renders EU citizens eligible for UB II receipt even then (BMAS 2009).

Immigrants' right to stay in Germany can be refused if an immigrant is eligible for meanstested public support. Special protection is granted to immigrants from signatory states of the European Convention on Social and Medical Assistance of 1953.<sup>6</sup> These immigrants can stay in Germany even if they receive welfare benefits (Classen 2012).

Prior studies show no difference in take-up behavior for natives and immigrants (see, e.g., Riphahn 2001, Frick and Groh-Samberg 2007, Bruckmeier and Wiemers 2012). However, BMAS (2009) points out that the expiration of UB I generates a substantially higher transition rate to UB II receipt among immigrant than native households: immigrant households and thus their needs are larger while their income and wealth are smaller than natives'. As the public debate about the reforms enhanced awareness of the new benefit program many observers expect a general increase in the propensity to take up benefits given eligibility (e.g., Bruckmeier and Wiemers 2012).

#### 3 Data

Using household data from the German Socio-Economic Panel Study (SOEP) (Wagner et al. 2007), we conduct separate analyses for natives and immigrants. We set the immigration status of the household according to the status of the household head. Since the number of immigrant households is small in East Germany, we consider only households in West Germany.<sup>7</sup>

<sup>&</sup>lt;sup>5</sup> Ethnic Germans are former German citizens or those belonging to the German people. After World War II, many migrated to West Germany and were granted German citizenship (Dietz 1999).

<sup>&</sup>lt;sup>6</sup> This covers immigrants from EU member states, Iceland, Norway, and—importantly—Turkey.

<sup>&</sup>lt;sup>7</sup> Other studies use similar sample selection criteria (e.g., Kogan 2004, Riphahn 2004).

We study welfare transitions before and after 2005, when the last step of the Hartz Reforms was implemented. The pre- and post-reform samples cover the years 2000 to 2004 and 2005 to 2010, respectively. Households are selected if they are part of the sample in 2000 or 2005, which define the initial states.<sup>8</sup> Recipients of unemployment assistance and UB II are required to be able to work. Therefore, we include only household heads of working age (25-60) and exclude the disabled.

Our dependent variable classifies households into three labor market states that indicate its status at the time of the survey. In a first step, we code all households who receive welfare benefits. Before the reform, we classify a household as receiving welfare if it receives one of the means-tested benefit schemes, i.e., social assistance or unemployment assistance. After the reform, we regard households as welfare recipients if the household receives UB II. To answer our research questions it would be sufficient to classify households as recipients vs. non-recipients. However, in order to refine the analysis and to derive additional insights, we split non-recipient households in two groups: non-recipient households are labelled "employed" if the household head is employed and "inactive" otherwise. The latter group includes household heads who are out of the labor force or unemployed and who may receive unemployment insurance benefits. The rationale for defining an "inactive group" is that these households neither work nor rely on welfare benefits but instead have other non-welfare income (e.g., unemployment insurance benefits or savings). It is possible that welfare receiving households have employed or unemployed heads; in both cases we code welfare receipt because earnings or unemployment benefits are topped up by welfare benefits.

Using weighted data to reflect the population of interest, Table 1 reports the observed annual distribution of the three labor market states for the pre-reform years 2000-2004 and the post-

<sup>&</sup>lt;sup>8</sup> The pre- and the post-reform samples cover periods of different length. We use fewer waves for the pre-reform period than for the post-reform period in order to be able to include the SOEP innovation sample F, which started in 2000

<sup>&</sup>lt;sup>9</sup> Across all years we observe that 17% and 30% of inactive native and immigrant households have unemployed heads, respectively.

reform years 2005-2010. In general, welfare recipient rates are higher after the reform with a noticeable jump shortly after the reform came into effect. The increase in recipiency rates is consistent with the decrease in non-take-up after the reform found by Bruckmeier and Wiemers (2012). After the reform, we observe rising employment and falling inactivity, reflecting the positive labor market trend and falling unemployment in this period.

There are remarkable differences between immigrants and natives. The share of immigrant households receiving welfare is more than twice as large as that of natives (e.g., in 2006: 15.2% vs. 7.1%). Correspondingly, the share of immigrant households that are classified as employed is considerably lower than that of natives.

Table 2 reports labor market transitions. Persistence is evident in all states. In the total population, welfare receipt has a persistence rate of more than 70%. Patterns change slightly from before to after the reform: while the welfare exit rate to employment increases (from 18.4% to 20.3%), welfare exit to inactivity becomes less frequent (from 10.3% to 6.1%). Labor market transitions appear to be less favorable for immigrants than for natives: immigrants have a much higher risk of welfare entry and less stable employment than natives. Welfare persistence increases for immigrants (from 68.1% to 75.4%) but not for natives, and welfare exit to employment increases for natives and declines for immigrants after the reform.

Table 3 shows descriptive statistics of the pre- and post-reform samples. Between the two sampling periods the share of female household heads increased and the mean number of children below age six in the households declined. The slight increases in mean age and education reflect changes in the population. Compared to natives, immigrant household heads have less education, are more often married, and have more children.<sup>10</sup>

<sup>&</sup>lt;sup>10</sup> In an online appendix we show further descriptive statistics by labor market and immigrant status: comparing welfare recipients and employed households, we observe small differences in the number of children while the share of married household heads is clearly smaller among welfare recipients. The share of single parents is considerably higher among welfare recipients than among households classified as inactive or employed. The figures also indicate a difference in average education between employed households and welfare recipients of two years among natives and one year among immigrants.

#### 4 Estimation strategy

The conceptual framework of our analysis follows the literature and uses a dynamic discrete choice model: a household chooses the labor market state (inactivity, employment, or welfare receipt) with the highest utility. Let  $U_{ijt}$  be the utility of household i in state j at time t:

$$U_{ijt} = \beta'_{i} \mathbf{x}_{it} + \gamma'_{i} \mathbf{y}_{i,t-1} + \alpha_{ij} + \varepsilon_{ijt}. \tag{1}$$

Utility depends on the observed household characteristics,  $\mathbf{x}_{it}$ .  $\boldsymbol{\beta}_j$  is a vector of alternative-specific coefficients. The coefficient vector  $\boldsymbol{\gamma}_j$  captures the effect of the previous state,  $\mathbf{y}_{i,t-1}$ , on the current state choice. We take account of household-specific unobserved heterogeneity by including a random error  $\alpha_{ij}$ .  $\varepsilon_{ijt}$  is an idiosyncratic error that is assumed to be independently distributed with a type I extreme value distribution.

Dynamic models of labor market state choice which allow for the presence of unobserved effects raise the problem of endogenous initial conditions: while transitions within the panel of observations are modeled, the transition to the very first observed state has no observed predecessor. We apply the conditional maximum likelihood estimator suggested by Wooldridge (2005) to solve this problem.

The specification of the Wooldridge approach models the unobserved heterogeneity  $\alpha_{ij}$  as a function of the initial state  $\mathbf{y}_{i0}$ , individual-specific explanatory variables  $\mathbf{x}_i$ , and a new random error,  $a_{ij}$ , that is uncorrelated with the initial state.<sup>11</sup> We assume  $a_{ij}$  to be normally distributed with zero mean and variance  $\sigma_a^2$ , i.e.,  $a_{ij}|(\mathbf{y}_{i0},\mathbf{x}_i) \sim N(0,\sigma_a^2)$ . Hence, the probability that individual i is in state j at time t conditional on observed and unobserved characteristics and the labor market

<sup>&</sup>lt;sup>11</sup> This approach in the spirit of Mundlak (1978) and follows the literature (see, e.g., Stewart 2007, Caliendo and Uhlendorff 2008, Cappellari and Jenkins 2009). Recently, Rabe-Hesketh and Skrondal (2013) discussed an alternative specification of the estimator to avoid potential biases. We present robustness tests along these suggestions in section 5.3 below.

state in t-1 can be written as

$$P(Y_{it} = j | \mathbf{x}_i, \mathbf{y}_{i,t-1}, \mathbf{y}_{i0}, \mathbf{a}_i) = \frac{\exp(\beta_j' \mathbf{x}_{it} + \gamma_j' \mathbf{y}_{i,t-1} + \delta_{j1}' \mathbf{y}_{i0} + \delta_{j2}' \mathbf{x}_i + a_{ij})}{\sum\limits_{k=1}^{\infty} \exp(\beta_k' \mathbf{x}_{it} + \gamma_k' \mathbf{y}_{i,t-1} + \delta_{k1}' \mathbf{y}_{i0} + \delta_{k2}' \mathbf{x}_i + a_{ik})}.$$
 (2)

Normalizing the coefficient vectors  $\beta_1, \gamma_1, \delta_{11}, \delta_{12}$ , and the unobserved heterogeneity,  $a_{i1}$ , to zero for the first alternative (k = 1), we can estimate a dynamic multinomial logit model with random effects.<sup>12</sup> We use Gauss-Hermite quadrature to integrate the random effect out of the corresponding log-likelihood.<sup>13</sup>

We use predicted probabilities for an individual randomly sampled from the population to describe state dependence in labor market transitions. To assess the incidence of true state dependence, we compare the predicted probability of welfare persistence to the predicted probability of entering welfare from inactivity. Our reasoning is as follows: if true state dependence exists, then the above two probabilities will differ. Previous welfare recipients will then have a higher probability to receive welfare than previously inactive non-recipients. If true state dependence does not exist or is unimportant, then these probabilities will not be significantly different.

The calculation of the predicted probabilites  $\overline{P}$  requires integrating over the distribution of the random effect (Skrondal and Rabe-Hesketh 2009):

$$\overline{P}(Y_t = j | \mathbf{y}_{t-1}, \mathbf{y}_0, \mathbf{x}^0) = \int \hat{P}(Y_t = j | \mathbf{y}_{t-1}, \mathbf{x}^0, \boldsymbol{\alpha}) h(\boldsymbol{\alpha} | \mathbf{x}, \mathbf{y}_0; \boldsymbol{\delta}) d\boldsymbol{\alpha}, \tag{3}$$

<sup>&</sup>lt;sup>12</sup> For contributions in the literature on welfare transitions applying the Wooldridge procedure, see Cappellari and Jenkins (2009) or Hansen and Lofstrom (2011). Erdem and Sun (2001) also apply this approach.

<sup>&</sup>lt;sup>13</sup> We use the Stata program -gllamm- written by Rabe-Hesketh et al. (2004).

<sup>&</sup>lt;sup>14</sup> We do not compare the probability of welfare persistence to the probability of entering welfare from employment because previously employed individuals are generally entitled to unemployment insurance benefits. Thus, they are expected to have a lower probability of entering welfare in the case of job loss.

where we set the vector  $\mathbf{x}^0$  to equal the sample average of the control variables.  $\hat{P}$  is the conditional probability. We assess the uncertainty of the prediction by approximate 95% confidence intervals for the predicted probabilities. <sup>15</sup>

#### 5 Results

This section presents the results obtained from dynamic multinomial logit models. To answer our first and second research question about state dependence in the German welfare system and its change over time, we describe patterns of welfare transitions and highlight changes in the dynamics after the reform in section 5.1. Section 5.2 turns to third question how welfare transitions relate to labor market conditions. The discussion addresses differences between immigrants and natives. We report results on robustness checks in section 5.3.

#### 5.1 Welfare transitions and state dependence

Table 4 shows the estimation results for the full sample. The positive estimates of the  $\gamma_j$  coefficients presented in the first rows indicate persistence in labor market states: employment in t-1 is associated with higher log-odds of employment in t and welfare receipt in t-1 is associated with higher log-odds of welfare receipt in t, both relative to inactivity.

Interestingly, welfare receipt in t-1 is linked to higher log-odds of employment in t relative to inactivity. This suggests that the welfare system incentivizes welfare recipients to take up employment. The size and significance of the coefficient of lagged employment as a determinant of welfare receipt in period t change between the pre- and post-reform periods. While the coefficient estimate is near zero (0.07) and statistically insignificant before the reform, it is larger (0.56) and

<sup>&</sup>lt;sup>15</sup> We use a parametric bootstrap approach with 1000 random draws from the sampling distribution of parameters. The procedure is available in the Stata ado-files -gllapred- and -ci\_marg\_mu- (Rabe-Hesketh et al. 2004, Skrondal and Rabe-Hesketh 2009).

statistically significant in the post-reform period. We return to the employment-to-welfare transition in greater detail below.

The control variables in Table 4 generally show similar patterns before and after the reform; e.g., higher education increases the log-odds of employment relative to inactivity and makes transitions to welfare relatively less likely. After the reform, the gender-specific life cycle patterns of labor market transitions are estimated more precisely. The individual-specific error term components significantly improve the model fit. The specification takes account of the potential endogeneity of health and the number of children (see variables labeled M in Table 4). The initial labor market state as of period t=0 yields highly significant coefficients, suggesting that the initial state indeed matters in explaining the current state.

Next, we discuss model-based predictions in Table 5 separately for the entire population (Panels A and B), native households (Panels C and D), and immigrant households (Panels E and F). <sup>16</sup> The calculations are based on separate estimations and average characteristics of the respective pre- and post-reform subsamples. <sup>17</sup> In part, these predictions clearly differ from the observed transitions. In particular, the predicted probabilities of welfare persistence are much lower compared to the observed probabilities, suggesting that characteristics explain welfare persistence to a large extent. <sup>18</sup>

The predictions reveal four interesting results. First, they confirm the persistence in labor market states in both sample periods. The probability of each labor market state in t is highest when the household was already in that state in t-1. Comparing pre and post periods, we observe a decline in the probability of welfare persistence, from 8.3% to 6.0%, i.e. by 28%.

<sup>&</sup>lt;sup>16</sup> The estimation results for natives and immigrants are presented in the online appendix.

<sup>&</sup>lt;sup>17</sup> For comparison, we also calculated predicted probabilities as the average of individually predicted probabilities. The results are similar in nature to those discussed and are presented in the online appendix.

<sup>&</sup>lt;sup>18</sup> While the observed probabilities of welfare persistence are about 73% for natives, the predicted probabilities are only about 6%. Other studies report similar results. For example, Hansen and Lofstrom (2009) observe that about 66% of natives stay on welfare in consecutive years. After controlling for initial conditions and heterogeneity, their predicted probability amounts only to about 10%.

Separate analyses for immigrants and natives show that the decline in welfare persistence is more pronounced among immigrants (Panels C-F).

Second, we find a clear increase in the probability of employment-to-welfare transitions, from 0.9% to 1.6% (Panels A and B). Judging from the non overlapping confidence intervals, the increase is statistically significant. The same pattern is observed for the immigrant and native samples. Although the overall risk of this transition is small, the sharp relative increase by 78 percent for the full sample is noteworthy, because we hardly observe other statistically significant changes over time. In addition, this result is remarkable because individuals are typically entitled to unemployment insurance benefits in the case of job loss (cf. section 2). Hence, one would expect that in the case of a job loss newly unemployed workers move from employment to the state of inactivity, which includes the receipt of unemployment insurance benefits. The increased risk of employment-to-welfare transitions may result from an increased propensity to take up short-term or low paid employment: short-term employment may be insufficient to generate UB I eligibility for the period after the short-term contract expired; low paid employment may not cover household needs and thus may go along with welfare benefit eligibility.

Third, the probability of a transition to employment increased after the reform; in addition, welfare-to-employment transitions are considerably more likely than inactivity-to-employment transitions. The increase in the probability of welfare exit to employment is particularly pronounced among immigrants (from 69% to 83%, Panels E and F). Among immigrants, the probability of inactivity-to-employment transition increased by approximately twenty percentage points, which is the largest absolute change. For both groups, persistence in inactivity declined after the reform. In general, this suggests that work incentives for welfare recipients and inactive households increased and that newly introduced activation measures might be effective, particularly among immigrants. In addition, this could indicate the availability of new job opportunities.

Next, we examine to what extent the changes in transition probabilities from before to after the reform are explained by changes in characteristics. For that purpose, we simulate post-reform transition probabilities for pre-reform characteristics, i.e., we calculate transition probabilities using household characteristics for the pre-reform period and the coefficients for the post-reform period. We argue that changes in transition probabilities can be attributed to changes in characteristics if the simulated probabilities converge to those originally predicted for the pre-reform period. Changes in labor market transitions may instead be attributed to changes in coefficients if the simulated probabilities converge to those originally predicted for the post-reform period.

Comparing the simulated probabilities and the originally predicted post-reform transitions, <sup>19</sup> we find almost identical results for most transitions, but some differences also emerge. On the one hand, we detect similar probabilities of transitions to employment. Thus, the substantial increases in transitions to employment, particularly among immigrants, are likely due to a change of the behavior of inactive or welfare-receiving households. On the other hand, the simulated probabilities differ from the original probabilities with respect to employment-to-welfare transitions and welfare persistence. The increase in welfare entry from employment is even more pronounced for pre-reform characteristics than for post-reform characteristics. Also, simulated welfare persistence is higher compared to the predictions using post-reform characteristics. This suggests that the change in characteristics even dampens the propensities to enter and to stay on welfare. Altogether, changes in transition patterns do not appear to be driven by changes in characteristics.

#### 5.2 Welfare transitions and labor market conditions

To address our third research question, this section investigates how welfare transitions relate to the labor market situation. Hoynes (2000) studies this relationship based on Californian administrative data. She confirms significant correlations between local labor markets, the duration of

<sup>&</sup>lt;sup>19</sup> The simulated probabilities are in Table A11 in the online appendix. The original post-reform transitions are in Table 5, Panels B, D, and F.

welfare receipt, and recidivism. We add state unemployment rates and their interactions with lagged labor market states to our specifications for the pre- and post-reform periods. This allows us to infer whether welfare transitions vary with labor market conditions.<sup>20</sup> Jointly, the three additional coefficients are statistically significant in three out of four models.

In general, the log-odds of employment decrease and the log-odds of welfare receipt increase with rising unemployment. To ease comparison between the pre- and post-reform periods, we present transition probabilities as a function of the unemployment rate graphically in Figures 1 und 2, after separate estimations for the native and immigrant subsamples.<sup>21</sup>

Among natives, state persistence hardly varies with the unemployment rate (Figure 1.1). Among immigrants, the curve for welfare persistence features a steeper slope after the reform (see dotted line in Figure 2.1). Thus, immigrant welfare persistence became more responsive to unemployment. Also, the downward shift of the curve indicates a general decrease in immigrants' welfare persistence.

For both natives and immigrants, the probability of welfare entry (from inactivity as well as from employment) increases with rising unemployment (Figures 1.2 and 2.2). This pattern hardly changed after the reform. Among immigrants, welfare entry from inactivity is less sensitive to the unemployment rate after than before the reform. As the overall probability of welfare entry declined, the reform incentives may have fostered additional job search activities.

Welfare exit to employment is less likely in periods of high unemployment (Figures 1.3 and 2.3). Overall, the rate of welfare-to-employment transitions increased after the reform; while the responsiveness of welfare exit to the unemployment rate hardly changed for natives it increased for immigrants: the dashed line in Figure 2.3 is considerably steeper after the reform.

In sum, welfare transitions are clearly correlated with labor market conditions. Immigrants' unemployment gradients of welfare persistence and welfare exit to employment are considerably

<sup>&</sup>lt;sup>20</sup> The parameter estimates for the pre- and post-reform period are available in the online appendix.

<sup>&</sup>lt;sup>21</sup> Due to the small number of observations, the predicted pre- and post-reform transition patterns are not significantly different and we omit the presentation of confidence intervals.

higher after than before the reform. This increased labor market responsiveness might reflect that immigrants benefit from the job creation in the economic boom early after the reform.

#### 5.3 Robustness checks

We perform five tests to check the robustness of our results. First, we address a potential measurement error in the initial condition. So far, we used the labor market state as of 2005, which typically was measured shortly after the reform. At that time, former recipients of unemployment insurance benefits might not have been aware of the precise institutional reforms that took effect on January 1, 2005 and they may have falsely indicated their benefit type. We omitted the 2005 data, started our window of observation in 2006 instead and re-estimated the model setting the initial condition to 2006. Based on predictions from these estimation results we find that the results are similar to those presented above. <sup>22</sup> In particular, trends in welfare entry and welfare exit are equivalent to those found in the full sample. Thus, our results are not driven by measurement error in the 2005 data.

Our second robustness check evaluates the relevance of the initial conditions control. We calculate predicted transition rates after setting the initial state to welfare receipt. Table 6 reports the results for individuals with average characteristics of welfare recipients. Again, the persistence of inactivity and welfare receipt declines for natives, however, it reaches higher levels than observed in Table 5. Among immigrants a decline in welfare persistence cannot be confirmed. Their welfare persistence again is not significantly higher than the probability of moving from inactivity to welfare. The initial condition controls explain a substantial part of the state dependence observed in the raw data (Table 2): the low welfare persistence for households with average characteristics in Table 5 is largely connected to the control for endogenous initial conditions.

<sup>&</sup>lt;sup>22</sup> The results are available in the online appendix.

Our third robustness check addresses the definition of the dependent variable. A feature of UB II is that it is paid to all households in need of support even if their members are employed (see section 2). Households with employed welfare recipients are called *Aufstocker*. In our definition of the dependent variable, Aufstocker are coded as welfare recipients. We re-estimate our model after coding Aufstocker as employed households, instead. Table 7 shows the predictions based on these estimations. We find the expected mechanical changes in transition probabilities: the transition rate from employment to welfare declines compared to Table 5 because households taking up welfare while employed no longer change their state. Also, the transition probability from inactivity to employment increases and that from inactivity to welfare decreases as a consequence of changed definitions. Our first key result, i.e., the decline in welfare persistence after the reform, no longer holds with redefined outcomes. Now, welfare persistence slightly increases after the reform for natives, and for immigrants we hardly observe a change in welfare persistence. Jointly, the results in Tables 5 and 7 suggest, that households who receive welfare while being employed are more likely to leave welfare dependence after than before the reform. This apparently drives the decline in welfare persistence in Table 5. The other two key results, i.e., the strong increase in the transition rate from employment to welfare and the increasing transition rates from inactivity and welfare to employment are generally confirmed with the recoded dependent variable.

As a fourth robustness check, we re-estimated our transition models controlling for federal state fixed effects. These fixed effect failed to be jointly statistically significant and the predicted transition patterns hardly differ from the overall pattern described in Table 5 above.<sup>23</sup>

Finally, we respond to the suggestion by Rabe-Hesketh and Skrondal (2013) and test whether our approach to the solution of the initial conditions problem biases the results. In addition to individual-specific averages of time-varying explanatory covariates, we also include initial-period explanatory variables as regressors in a more flexible specification.<sup>24</sup> We find that the relevant

<sup>&</sup>lt;sup>23</sup> The results are available upon request.

<sup>&</sup>lt;sup>24</sup> Tables A12 and A13 in the online appendix show the estimation results and the corresponding transition matrices, respectively.

results from the extended specification, i.e., the estimated coefficients of the lagged variables and the predicted probabilities of labor market transitions, differ from those of the constrained specification by less than two percent. We regard this as persuasive evidence that our specification is not biased due to an overly restrictive model.

#### 6 Conclusion

We use dynamic multinomial logit models to analyze welfare transitions and to determine the role of state dependence in the German welfare system. We study welfare dynamics before and after a reform and explore the relevance of labor market conditions for welfare transitions. We compare welfare transitions of immigrants and natives accounting for the endogeneity of initial conditions and unobserved heterogeneity.

We draw four main conclusions: first, true state dependence in welfare receipt is not a dominant factor for welfare receipt in Germany. The probability of welfare persistence is not significantly higher than the probability of entering welfare from inactivity. Second, our evidence suggests that the pre- and post-reform transition patterns differ. In particular, the transition to employment became more likely and the persistence in welfare receipt and inactivity declined. This may suggest that the reform enhanced labor market attachment and work incentives for welfare recipients and inactive individuals. Third, immigrants' responsiveness to the labor market has increased after the reform, e.g., with respect to welfare persistence and welfare exit. Finally, the overall decline in welfare persistence after the reform seems to be due to those households who receive welfare to top up their labor market earnings. After the reform, this group has a higher propensity to leave welfare receipt than prior to the reform.

Our evidence shows that the labor market situation contributes to explain welfare transitions, i.e., welfare entry is lower and welfare exit is higher when unemployment is low. This finding agrees with the results of Hoynes (2000). Our analysis also points to a change after the reform

that may not have been intended: there is a substantial increase in the employment-to-welfare transition rate, i.e., the rate at which households start to receive welfare given that the head of the household was employed before. Several explanations are plausible: first, households might have become more likely to fall below the eligibility threshold despite employment if they earn lower incomes than before. Second, employed households may not accumulate sufficient claims for unemployment insurance benefits (UB I) after the reform if their employment spells are shorther than the required minimum contribution period to the unemployment insurance. In that case a loss of employment can generate welfare dependence because a claim against the unemployment insurance for UB I is not established.<sup>25</sup>

Overall, true state dependence is not an important factor in the German welfare system even though its level certainly varies across population groups. The patterns and dynamics of welfare transitions changed from before to after the reform in a way that is consistent with the reform objectives: after the reform, non-working households display higher labor market attachments, and welfare transitions are more responsive to the labor market situation. The reform of the German welfare system may be instructive for other countries that intend to promote work incentives in the presence of troubled labor markets. It seems to be feasible to provide a safety net without a welfare trap.

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<sup>&</sup>lt;sup>25</sup> Jahn and Stephan (2012) show that about 18% of those who became unemployed in 2010 moved directly into UB II instead of UB I. Koller and Rudolph (2011) describe that welfare exit after the Hartz Reforms generated unstable employment situations, as only 55% of the new jobs last longer than six months.

#### References

- Barrett, A. and McCarthy, Y. (2008). Immigrants and welfare programmes: exploring the interactions between immigrant characteristics, immigrant welfare dependence, and welfare policy. *Oxford Review of Economic Policy*, 24(3):543–560.
- Bitler, M. P. and Hoynes, H. W. (2010). The state of social safety net in the post-welfare reform era. *Brookings Papers on Economic Activity*, 41(2):71–147.
- Blank, R. M. (2002). Evaluating welfare reform in the United States. *Journal of Economic Literature*, 40(4):1105–1166.
- BMAS (Bundesministerium für Arbeit und Soziales). Wirkungen des SGB II auf Personen mit Migrationshintergrund. Abschlussbericht [online]. (2009). Available from: http://www.bmas. de/portal/39960/. Last accessed: Sept. 13, 2013.
- Brewer, M., Duncan, A., Shephard, A., and Suarez, M. J. (2006). Did working families' tax credit work? The impact of in-work support on labour supply in Great Britain. *Labour Economics*, 13(6):699–720.
- Bruckmeier, K. and Wiemers, J. (2012). A new targeting: a new take-up? Non-take-up of social assistance in Germany after social policy reforms. *Empirical Economics*, 43(2):565–580.
- Burda, M. C. and Hunt, J. (2011). What explains the German labor market miracle in the great recession. *Brookings Papers on Economic Activity*, 42(1):273–335.
- Caliendo, M. and Hogenacker, J. (2012). The German labor market after the Great Recession: successful reforms and future challenges. *IZA Journal of European Labor Studies*, 1(1):3.
- Caliendo, M. and Uhlendorff, A. (2008). Self-employment dynamics, state dependence and cross-mobility patterns. IZA Discussion Papers 3900, Institute for the Study of Labor (IZA), Bonn.
- Cappellari, L. and Jenkins, S. P. (2008). The dynamics of social assistance receipt: measurement and modelling issues, with an application to Britain. OECD Social, Employment and Migration Working Papers 67, OECD, Directorate for Employment, Labour and Social Affairs.
- Cappellari, L. and Jenkins, S. P. (2009). The dynamics of social assistance benefit receipt in Britain. IZA Discussion Papers 4457, Institute for the Study of Labor (IZA), Bonn.
- Chay, K. Y., Hoynes, H. W., and Hyslop, D. (2004). True state dependence in monthly welfare participation: a nonexperimental analysis. Working Papers 05-33, University of California at Davis, Department of Economics.
- Classen, G. Sozialleistungen für MigrantInnen nach SGB II, SGB XII und AsylbLG [online]. (2012). Available from: http://www.fluechtlingsinfo-berlin.de/fr/pdf/Classen\_SGB\_II\_XII\_AsylbLG.pdf. Last accessed: Nov. 16, 2012.
- Dietz, B. (1999). Ethnic German immigration from Eastern Europe and the former Soviet Union to Germany: the effects of migrant networks. IZA Discussion Papers 68, Institute for the Study of Labor (IZA), Bonn.
- Dietz, M., Koch, S., Rudolph, H., Walwei, U., and Wiemers, J. (2011). Reform der Hinzuverdienstregeln im SGB II. Fiskalische Effekte und Arbeitsmarktwirkungen. *Sozialer Fortschritt*, 60(1-2):4–15.

- Edmark, K. (2009). Migration effects of welfare benefit reform. *Scandinavian Journal of Economics*, 111(3):511–526.
- Erdem, T. and Sun, B. (2001). Testing for choice dynamics in panel data. *Journal of Business and Economic Statistics*, 19(2):142–152.
- Fahr, R. and Sunde, U. (2009). Did the Hartz reforms speed-up the matching process? A macroevaluation using empirical matching functions. *German Economic Review*, 10(3):284–316.
- Fortin, B., Lacroix, G., and Drolet, S. (2004). Welfare benefits and the duration of welfare spells: evidence from a natural experiment in Canada. *Journal of Public Economics*, 88(7-8):1495–1520.
- Frick, J. R. and Groh-Samberg, O. (2007). To claim or not to claim: estimating non-take-up of social assistance in Germany and the role of measurement error. Discussion Papers of DIW Berlin 734, DIW Berlin, German Institute for Economic Research.
- Grogger, J. and Karoly, L. A. (2005). *Welfare Reform: Effects of a Decade of Change*. Harvard University Press, Cambridge, Massachusetts.
- Hansen, J. and Lofstrom, M. (2009). The dynamics of immigrant welfare and labor market behavior. *Journal of Population Economics*, 22(4):941–970.
- Hansen, J. and Lofstrom, M. (2011). Immigrant-native differences in welfare participation: the role of entry and exit rates. *Industrial Relations: A Journal of Economy and Society*, 50(3):412–442.
- Hansen, J., Lofstrom, M., and Zhang, X. (2006). State dependence in Canadian welfare participation. IZA Discussion Papers 2266, Institute for the Study of Labor (IZA), Bonn.
- Hoynes, H. W. (2000). Local labor markets and welfare spells: Do demand conditions matter? *The Review of Economics and Statistics*, 82(3):351–368.
- Hoynes, H. W., Miller, D. L., and Schaller, J. (2012). Who suffers during recessions? *Journal of Economic Perspectives*, 26(3):27–48.
- Huber, M., Lechner, M., Wunsch, C., and Walter, T. (2011). Do German welfare-to-work programmes reduce welfare dependency and increase employment? *German Economic Review*, 12(2):182–204.
- Jahn, E. and Stephan, G. (2012). Arbeitslosengeld wie lange man dafür arbeiten muss. IAB Kurzbericht 19, IAB, Nürnberg.
- Klinger, S. and Rothe, T. (2012). Der Rückgang der Langzeitarbeitslosigkeit in Deutschland: Ein Erfolg der Hartz-Reformen oder konjunktureller Effekt? *Schmollers Jahrbuch (Journal of Applied Social Science Studies)*, 132(1):1–33.
- Kogan, I. (2004). Last hired, first fired? The unemployment dynamics of male immigrants in Germany. *European Sociological Review*, 20(5):445–461.
- Koller, L. and Rudolph, H. (2011). Viele Jobs von kurzer Dauer. IAB Kurzbericht 14, IAB, Nürnberg.
- Mundlak, Y. (1978). On the pooling of time series and cross section data. *Econometrica*, 46(1):69–85.

- Rabe-Hesketh, S. and Skrondal, A. (2013). Avoiding biased versions of Wooldridge's simple solution to the initial conditions problem. *Economics Letters*, 120(2):346–349.
- Rabe-Hesketh, S., Skrondal, A., and Pickles, A. (2004). Gllamm manual. U.C. Berkeley Division of Biostatistics Working Paper Series 1160, Berkeley Electronic Press.
- Riphahn, R. T. (2001). Rational poverty or poor rationality? The take-up study of social assistance benefits. *Review of Income and Wealth*, 47(3):379–98.
- Riphahn, R. T. (2004). Immigrant participation in social assistance programs: evidence from German guestworkers. *Applied Economics Quarterly*, 50(4):329–362.
- Riphahn, R. T. and Wunder, C. (2013). Patterns of welfare dependence before and after a reform: Evidence from first generation immigrants and natives in Germany. *Review of Income and Wealth*, 59(3):437–459.
- Schneider, H. (2012). Wie nachhaltig ist das deutsche Jobwunder? Eine Reformbilanz. IZA Standpunkte 51, Institute for the Study of Labor (IZA), Bonn.
- Skrondal, A. and Rabe-Hesketh, S. (2009). Prediction in multilevel generalized linear models. *Journal Of The Royal Statistical Society Series A*, 172(3):659–687.
- Stewart, M. B. (2007). The interrelated dynamics of unemployment and low-wage employment. *Journal of Applied Econometrics*, 22(3):511–531.
- SVR (Sachverständigenrat zur Begutachtung der gesamtwirtschaftlichen Entwicklung) (2011). *Verantwortung für Europa wahrnehmen.* Jahresgutachten 2011/12. Wiesbaden.
- SVR (Sachverständigenrat zur Begutachtung der gesamtwirtschaftlichen Entwicklung). Zeitreihen für Deutschland [online]. (2012). Available from: http://www.sachverstaendigenrat-wirtschaft.de/zr\_deutschland.html. Last accessed: Nov. 16, 2012.
- Wagner, G. G., Frick, J. R., and Schupp, J. (2007). The German Socio-Economic Panel Study (SOEP) scope, evolution and enhancemants. *Schmollers Jahrbuch (Journal of Applied Social Science Studies)*, 127(1):139–169.
- Wooldridge, J. M. (2005). Simple solutions to the initial conditions problem in dynamic, non-linear panel data models with unobserved heterogeneity. *Journal of Applied Econometrics*, 20(1):39–54.
- Ziliak, J. P., Figlio, D. N., Davis, E. E., and Connolly, L. S. (2000). Accounting for the decline in AFDC caseloads: welfare reform or the economy? *Journal of Human Resources*, 35(3):570–586.

# **Tables and Figures**

Table 1 Observed distribution of labor market states by year

Year		State at time <i>t</i>		Sample size
	Inactivity	Employment	Welfare	_
A. Total population:	pre-reform			
2000	10.44	85.75	3.81	5,082
2001	9.86	86.21	3.93	4,871
2002	11.52	83.73	4.75	4,268
2003	12.17	82.34	5.49	3,951
2004	11.12	82.27	6.60	3,644
Total	10.97	84.2	4.83	21,816
B. Total population:	post-reform			
2005	13.14	80.72	6.14	3,873
2006	11.52	80.21	8.27	3,736
2007	10.59	82.33	7.08	3,359
2008	8.87	84.70	6.43	3,057
2009	9.09	85.45	5.46	2,698
2010	8.87	84.04	7.09	2,401
Total	10.51	82.74	6.75	19,124
C. Natives: pre-refo				17,12.
2000	9.95	86.96	3.09	4,163
2001	9.17	87.64	3.19	3,997
2002	11.09	85.16	3.75	3,510
2003	11.56	83.72	4.73	3,258
2004	10.94	83.25	5.82	3,016
Total	10.47	85.49	4.03	17,944
		03.47	4.03	17,344
D. Natives: post-ref 2005		01 06	5 57	2 260
	12.57	81.86	5.57	3,260
2006	11.40	81.48	7.12	3,145
2007	10.61	83.43	5.96	2,845
2008	8.54	85.70	5.76	2,603
2009	9.17	86.31	4.52	2,313
2010	8.63	85.46	5.92	2,071
Total	10.30	83.87	5.83	16,237
E. Immigrants: pre-				
2000	13.73	77.55	8.71	919
2001	14.55	76.49	8.95	874
2002	14.22	74.75	11.03	758
2003	16.04	73.67	10.29	693
2004	12.32	75.99	11.68	628
Total	14.21	75.72	10.07	3,872
F. Immigrants: post-	-reform			
2005	17.38	72.23	10.40	613
2006	12.26	72.56	15.18	591
2007	10.49	75.38	14.13	514
2008	11.02	78.21	10.77	454
2009	8.57	79.82	11.60	385
2010	10.50	74.38	15.13	330
Total	11.86	75.27	12.87	2,887

Note: Percentage of households weighted using cross-sectional weights.

Source: SOEP 2001-2010.

Table 2 Observed probabilities of labor market transitions

State in $t-1$	State at time <i>t</i>						
_	Inactivity	Employment	Welfare				
A. Total population: pre-	-reform						
Inactivity	0.633	0.272	0.095				
Employment	0.047	0.944	0.010				
Welfare receipt	0.103	0.184	0.712				
B. Total population: pos	t-reform						
Inactivity	0.607	0.299	0.095				
Employment	0.037	0.950	0.013				
Welfare receipt	0.061	0.203	0.736				
C. Natives: pre-reform							
Inactivity	0.645	0.281	0.075				
Employment	0.043	0.949	0.008				
Welfare receipt	0.104	0.168	0.728				
D. Natives: post-reform							
Inactivity	0.616	0.305	0.079				
Employment	0.034	0.954	0.011				
Welfare receipt	0.069	0.204	0.727				
E. Immigrants: pre-refor	rm						
Inactivity	0.590	0.242	0.169				
Employment	0.066	0.915	0.019				
Welfare receipt	0.103	0.216	0.681				
F. Immigrants: post-refo	rm						
Inactivity	0.576	0.276	0.149				
Employment	0.050	0.928	0.023				
Welfare receipt	0.045	0.201	0.754				

*Note*: Share of household heads weighted using cross-sectional weights. *Source*: SOEP 2000-2010.

Table 3
Descriptive statistics

		Pre-reform	(2000-2004)		Post-reform (2005-2010)				
	Nativ	es	Immig	grants	Nati	ves	Immigrants		
Variable	Mean	Std. Dev.	Mean	Std. Dev.	Mean	Std. Dev.	Mean	Std. Dev.	
Inactivity	0.11	0.31	0.15	0.35	0.10	0.30	0.11	0.32	
Employment	0.86	0.35	0.77	0.42	0.86	0.35	0.78	0.42	
Welfare	0.04	0.18	0.09	0.28	0.05	0.21	0.11	0.31	
Age	43.09	8.57	42.52	9.17	44.25	8.39	43.57	8.74	
Female	0.35	0.48	0.27	0.44	0.41	0.49	0.37	0.48	
Education in years	12.59	2.74	11.00	2.41	12.75	2.75	11.31	2.52	
Married	0.66	0.47	0.79	0.40	0.63	0.48	0.78	0.42	
Health status: good	0.60	0.49	0.57	0.50	0.55	0.50	0.54	0.50	
School in Germany: no	0.00	0.00	0.60	0.49	0.00	0.00	0.47	0.50	
Number of children LT6	0.23	0.52	0.33	0.60	0.17	0.45	0.24	0.52	
Number of children GE6	0.57	0.86	0.81	0.99	0.52	0.83	0.80	0.96	
Year 2001	0.29	0.45	0.30	0.46	0.00	0.00	0.00	0.00	
Year 2002	0.26	0.44	0.26	0.44	0.00	0.00	0.00	0.00	
Year 2003	0.24	0.43	0.24	0.42	0.00	0.00	0.00	0.00	
Year 2004	0.22	0.41	0.21	0.41	0.00	0.00	0.00	0.00	
Year 2006	0.00	0.00	0.00	0.00	0.24	0.43	0.26	0.44	
Year 2007	0.00	0.00	0.00	0.00	0.22	0.41	0.23	0.42	
Year 2008	0.00	0.00	0.00	0.00	0.20	0.40	0.20	0.40	
Year 2009	0.00	0.00	0.00	0.00	0.18	0.38	0.17	0.38	
Year 2010	0.00	0.00	0.00	0.00	0.16	0.37	0.15	0.35	
Initial condition (in 2005)									
Inactivity	0.10	0.30	0.13	0.34	0.11	0.32	0.17	0.38	
Employment	0.87	0.34	0.78	0.41	0.85	0.36	0.74	0.44	
Welfare receipt	0.03	0.18	0.09	0.28	0.04	0.19	0.09	0.29	
Number of person-year observations	13,781	1	2,953	3	12,97	77	2,274	1	

Source: SOEP 2000-2010.

Table 4
Estimation results: total population

Variable		Pre refe	om		Post-reform			
	Employme	Employment		Welfare receipt		ent	Welfare receipt	
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.
Employed in t-1	2.276***	(0.142)	0.071	(0.220)	2.182***	(0.141)	0.561**	(0.223)
Welfare receipt in t-1	1.415***	(0.242)	2.041***	(0.281)	1.485***	(0.238)	1.789***	(0.269)
Age	0.451***	(0.075)	0.162	(0.124)	0.738***	(0.089)	0.550***	(0.153)
Age squared	-0.565***	(0.083)	-0.196	(0.137)	-0.884***	(0.098)	-0.646***	(0.169)
Female	0.535	(2.076)	3.033	(3.408)	7.965***	(2.356)	13.590***	(4.026)
$Age \times Female$	-0.118	(0.097)	-0.155	(0.158)	-0.492***	(0.110)	-0.672***	(0.187)
Age sq. × Female	0.170	(0.110)	0.155	(0.178)	0.610***	(0.124)	0.753***	(0.210)
Education	0.113***	(0.019)	-0.177***	(0.036)	0.055***	(0.021)	-0.178***	(0.041)
School in Germany: no	-0.421***	(0.144)	0.597***	(0.206)	-0.016	(0.200)	0.616**	(0.294)
Married	-0.223**	(0.110)	-1.123***	(0.177)	-0.644***	(0.125)	-1.673***	(0.207)
Health status: good	0.018	(0.105)	-0.473***	(0.173)	-0.158	(0.117)	-0.612***	(0.187)
No. of kids LT 6	0.542***	(0.134)	0.659***	(0.235)	0.381***	(0.146)	0.186	(0.246)
No. of kids GE 6	0.240**	(0.120)	0.416**	(0.186)	0.160	(0.132)	0.218	(0.204)
Year 2002	-0.143	(0.098)	0.061	(0.170)	_		_	
Year 2003	-0.322***	(0.100)	0.153	(0.173)	_		_	
Year 2004	-0.136	(0.106)	0.586***	(0.178)	_		_	
Year 2007	_		_		0.162	(0.115)	-0.385**	(0.181)
Year 2008	_		_		0.355***	(0.122)	-0.387**	(0.195)
Year 2009	_		_		0.229*	(0.127)	-0.674***	(0.212)
Year 2010	_		_		0.266**	(0.133)	-0.067	(0.212)
Employed in t=0	2.301***	(0.227)	0.239	(0.300)	2.576***	(0.222)	-0.382	(0.313)
Welfare receipt in t=0	-0.022	(0.293)	2.295***	(0.396)	0.402	(0.293)	2.991***	(0.410)
M: Health status: good	0.500***	(0.176)	-0.401	(0.293)	0.696***	(0.206)	-0.485	(0.349)
M: No. of kids LT 6	-1.373***	(0.202)	-0.799**	(0.365)	-1.166***	(0.265)	0.040	(0.421)
M: No. of kids GE 6	-0.227	(0.156)	-0.250	(0.241)	0.182	(0.145)	0.194	(0.230)
Constant	-9.783***	(1.623)	-2.826	(2.735)	-15.176***	(1.924)	-9.890***	(3.348)
$\overline{Var(a_{ij})}$	2.394	(0.380)	1.914	(0.563)	2.440	(0.365)	4.203	(0.847)
$Cov(a_{i,empl}, a_{i,welf})$	0.092	(0.399)			0.036	(0.405)		
log likelihood		-4936.09	963			-4317.50	091	
No. of household-year observations		16,73	4			15,25		
No. of households		5,094				3,882		

*Note*: Dynamic multinomial logit models with random effects. Dependent variable: labor market state (inactivity, employment, welfare receipt). M: denotes individual-specific averages of a variable. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01.

Source: SOEP 2000-2010.

Table 5
Predicted probabilities of labor market transitions
given subsample-period-specific average characteristics

State at time $t-1$	State at time <i>t</i>								
	Inactive		Е	Employment			Welfare		
	Mean	95%	ó-CI	Mean	95%	ó-CI	Mean	95%	ó-CI
A. Pre-reform									
Inactive	0.246	0.206	0.292	0.718	0.667	0.757	0.037	0.027	0.053
Employment	0.055	0.049	0.061	0.936	0.929	0.942	0.009	0.008	0.012
Welfare	0.089	0.064	0.126	0.828	0.762	0.868	0.083	0.056	0.133
B. Post-reform									
Inactive	0.195	0.161	0.237	0.767	0.724	0.800	0.038	0.029	0.053
Employment	0.042	0.036	0.048	0.942	0.935	0.948	0.016	0.014	0.021
Welfare	0.066	0.046	0.095	0.874	0.835	0.901	0.060	0.045	0.085
C. Natives: pre-reform									
Inactive	0.220	0.180	0.269	0.759	0.705	0.797	0.021	0.014	0.036
Employment	0.050	0.045	0.057	0.943	0.936	0.949	0.007	0.005	0.009
Welfare	0.081	0.053	0.120	0.854	0.787	0.894	0.065	0.040	0.117
D. Natives: post-reform	n								
Inactive	0.195	0.159	0.242	0.773	0.726	0.812	0.032	0.023	0.048
Employment	0.041	0.036	0.047	0.947	0.939	0.953	0.013	0.010	0.017
Welfare	0.067	0.045	0.100	0.877	0.829	0.907	0.057	0.039	0.089
E. Immigrants: pre-refe	orm								
Inactive	0.362	0.262	0.494	0.521	0.381	0.622	0.118	0.080	0.206
Employment	0.074	0.059	0.095	0.900	0.874	0.916	0.026	0.019	0.045
Welfare	0.123	0.070	0.201	0.693	0.513	0.781	0.184	0.119	0.362
F. Immigrants: post-ref	orm								
Inactive	0.204	0.127	0.326	0.721	0.591	0.793	0.076	0.047	0.134
Employment	0.046	0.033	0.064	0.914	0.885	0.932	0.040	0.028	0.064
Welfare	0.056	0.027	0.105	0.830	0.735	0.884	0.115	0.074	0.188

*Note*: Calculations are based on separate estimations for all subsamples in both periods. Estimation results are presented in the online appendix. Simulation-based 95% confidence intervals are calculated using 1000 replications.

Table 6
Predicted probabilities of labor market transitions
given subsample-period-specific average characteristics of welfare recipients
setting initial state to welfare

State at time $t-1$	State at time <i>t</i>									
	Inactive			Е	Employment			Welfare		
	Mean	95%	ó-CI	Mean	95%	5-CI	Mean	95%	ó-CI	
A. Natives: pre-reform										
Inactive	0.374	0.270	0.505	0.207	0.140	0.284	0.419	0.282	0.539	
Employment	0.190	0.126	0.284	0.509	0.401	0.617	0.301	0.187	0.415	
Welfare	0.095	0.064	0.137	0.190	0.135	0.251	0.715	0.646	0.776	
B. Natives: post-reform	n									
Inactive	0.236	0.162	0.328	0.251	0.182	0.329	0.513	0.401	0.611	
Employment	0.092	0.058	0.140	0.517	0.421	0.623	0.392	0.283	0.490	
Welfare	0.070	0.045	0.102	0.270	0.203	0.337	0.660	0.592	0.731	
C. Immigrants: pre-ref	orm									
Inactive	0.386	0.256	0.555	0.149	0.079	0.239	0.465	0.292	0.600	
Employment	0.186	0.103	0.294	0.570	0.406	0.708	0.245	0.129	0.396	
Welfare	0.133	0.089	0.200	0.218	0.145	0.292	0.649	0.559	0.740	
D. Immigrants: post-re	eform									
Inactive	0.244	0.134	0.415	0.221	0.126	0.333	0.535	0.383	0.662	
Employment	0.093	0.041	0.181	0.451	0.330	0.605	0.456	0.295	0.578	
Welfare	0.063	0.033	0.119	0.261	0.168	0.358	0.676	0.570	0.772	

*Note*: Calculations are based on separate estimations for all subsamples in both periods. Estimation results are presented in the online appendix. Simulation-based 95% confidence intervals are calculated using 1000 replications.

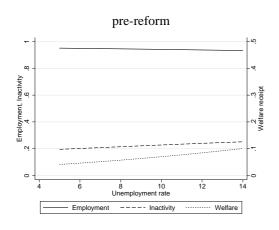
Table 7
Predicted probabilities of labor market transitions
given subsample-period-specific average characteristics (alternative definition of states)

State at time $t-1$	State at time <i>t</i>									
		Inactive		E	Employment			Welfare		
	Mean	Mean 95%-CI		Mean	95%	5-CI	Mean	95%	5-CI	
A. Total population: p	re-reform									
Inactive	0.253	0.211	0.303	0.723	0.669	0.764	0.025	0.017	0.037	
Employment	0.053	0.048	0.059	0.943	0.936	0.948	0.005	0.004	0.007	
Welfare	0.104	0.073	0.146	0.850	0.795	0.886	0.047	0.030	0.081	
B. Total population: po	ost-reform									
Inactive	0.194	0.160	0.236	0.782	0.739	0.816	0.024	0.017	0.037	
Employment	0.040	0.035	0.045	0.953	0.947	0.958	0.008	0.006	0.011	
Welfare	0.095	0.066	0.134	0.853	0.801	0.887	0.052	0.036	0.081	
C. Natives: pre-reform	l									
Inactive	0.224	0.183	0.275	0.765	0.712	0.804	0.011	0.007	0.020	
Employment	0.049	0.043	0.056	0.948	0.941	0.954	0.003	0.002	0.005	
Welfare	0.090	0.059	0.137	0.884	0.829	0.920	0.025	0.014	0.051	
D. Natives: post-reform	m									
Inactive	0.195	0.158	0.239	0.791	0.745	0.827	0.014	0.009	0.025	
Employment	0.039	0.034	0.045	0.955	0.948	0.960	0.006	0.005	0.009	
Welfare	0.098	0.061	0.149	0.866	0.803	0.909	0.036	0.022	0.065	
E. Immigrants: pre-ref	orm									
Inactive	0.391	0.280	0.535	0.504	0.354	0.615	0.105	0.068	0.209	
Employment	0.069	0.056	0.089	0.918	0.892	0.931	0.013	0.009	0.029	
Welfare	0.156	0.090	0.256	0.708	0.522	0.800	0.136	0.080	0.299	
F. Immigrants: post-re	form									
Inactive	0.201	0.132	0.316	0.722	0.584	0.802	0.077	0.045	0.153	
Employment	0.042	0.030	0.058	0.938	0.915	0.951	0.021	0.014	0.034	
Welfare	0.071	0.033	0.131	0.796	0.661	0.867	0.134	0.079	0.256	

*Note*: Calculations are based on separate estimations for all subsamples in both periods. Estimation results are presented in the online appendix. Simulation-based 95% confidence intervals are calculated using 1000 replications.

## Figure 1 Labor market transitions and unemployment rate (natives)

Fig. 1.1: Persistence in employment and welfare participation



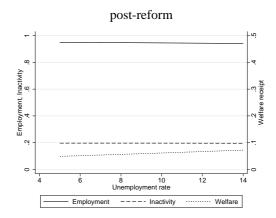
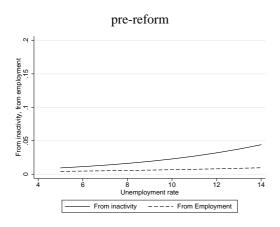


Fig. 1.2: Welfare entry



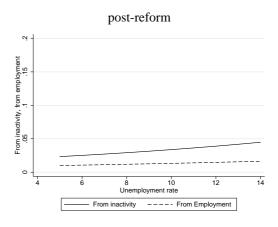
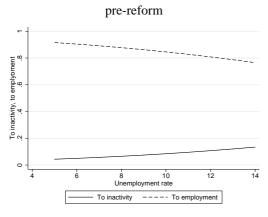
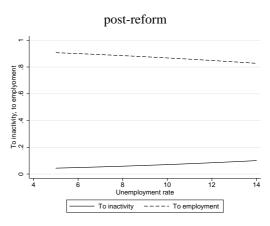


Fig. 1.3: Welfare exit

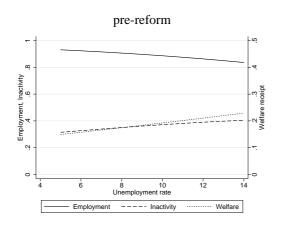




*Note*: Predicted probabilities given average characteristics. Figures 1.1 uses a secondary vertical axes to indicate transition probabilities.

## Figure 2 Labor market transitions and unemployment rate (immigrants)

Fig. 2.1: Persistence in employment and welfare participation



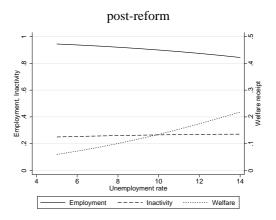
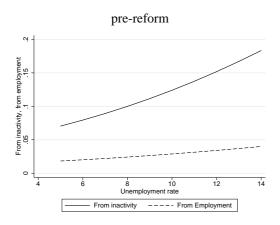


Fig. 2.2: Welfare entry



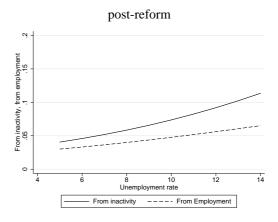
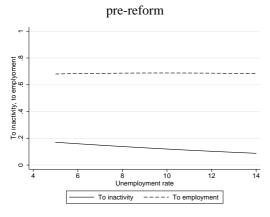
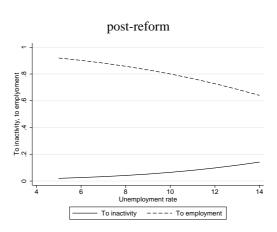


Fig. 2.3: Welfare exit





*Note*: Predicted probabilities given average characteristics. Figures 2.1 uses a secondary vertical axes to indicate transition probabilities.

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Table A1 Averages of selected variables by labor market state

	Pre-reform				Post-reform	
Variable	Inactivity	Employment	Welfare	Inactivity	Employment	Welfare
A. Natives						
Age	43.91	43.01	42.61	44.00	44.30	43.87
Female	0.65	0.31	0.65	0.69	0.37	0.64
Education in years	12.07	12.74	10.63	12.45	12.87	11.08
Married	0.70	0.67	0.36	0.72	0.63	0.32
Health status: good	0.56	0.62	0.39	0.54	0.56	0.33
School in Germany: no	0.00	0.00	0.00	0.00	0.00	0.00
Number of children LT6	0.36	0.21	0.31	0.34	0.15	0.22
Number of children GE6	0.49	0.57	0.68	0.50	0.52	0.61
Household size	2.92	2.82	2.70	2.98	2.72	2.48
Single person	0.16	0.19	0.22	0.13	0.20	0.28
Single parent	0.07	0.06	0.34	0.07	0.07	0.30
Couple Without Children	0.21	0.21	0.12	0.21	0.23	0.10
Couple with children	0.54	0.52	0.28	0.57	0.49	0.30
Other household type	0.02	0.02	0.03	0.02	0.01	0.02
B. Immigrants						
Age	42.56	42.18	45.42	44.63	43.49	43.10
Female	0.53	0.22	0.29	0.66	0.31	0.55
Education in years	10.52	11.17	10.40	10.79	11.52	10.41
Married	0.81	0.79	0.78	0.82	0.79	0.62
Health status: good	0.51	0.61	0.37	0.47	0.57	0.39
School in Germany: no	0.55	0.59	0.73	0.48	0.47	0.50
Number of children LT6	0.46	0.29	0.40	0.39	0.22	0.21
Number of children GE6	0.68	0.82	0.92	0.72	0.78	1.00
Household size	3.44	3.37	3.66	3.40	3.30	3.23
Single person	0.07	0.10	0.09	0.07	0.08	0.17
Single parent	0.09	0.07	0.12	0.08	0.09	0.27
Couple Without Children	0.17	0.15	0.15	0.16	0.16	0.06
Couple with children	0.63	0.67	0.55	0.68	0.65	0.50
Other household type	0.04	0.02	0.09	0.02	0.02	0.01

Source: SOEP 2001-2004 and 2006-2010.

Table A2
Separate estimation results for natives and immigrants: pre-reform

Variable		Native	es		Immigrants				
	Employme	ent	Welfare rec	eipt	Employme	ent	Welfare rec	eipt	
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	
Employed in t-1	2.218***	(0.161)	0.302	(0.268)	2.563***	(0.299)	-0.238	(0.390)	
Welfare receipt in t-1	1.387***	(0.307)	2.332***	(0.349)	1.548***	(0.399)	1.662***	(0.470)	
Age	0.524***	(0.090)	0.297*	(0.161)	0.259**	(0.131)	-0.015	(0.198)	
Age squared	-0.649***	(0.100)	-0.359**	(0.178)	-0.344**	(0.147)	0.040	(0.218)	
Female	2.151	(2.467)	4.574	(4.261)	-3.645	(3.966)	0.805	(6.039)	
Age × Female	-0.193*	(0.115)	-0.226	(0.197)	0.073	(0.188)	-0.051	(0.284)	
Age sq. $\times$ Female	0.252*	(0.129)	0.234	(0.221)	-0.033	(0.216)	0.041	(0.322)	
Education	0.117***	(0.022)	-0.255***	(0.048)	0.106***	(0.039)	0.008	(0.057)	
School in Germany: no	_		_		-0.292	(0.219)	0.330	(0.353)	
Married	-0.251**	(0.127)	-1.214***	(0.209)	-0.117	(0.238)	-0.488	(0.357)	
Health status: good	0.040	(0.123)	-0.308	(0.213)	-0.045	(0.208)	-0.809***	(0.302)	
No. of kids LT 6	0.684***	(0.157)	0.720**	(0.308)	0.141	(0.260)	0.709*	(0.375)	
No. of kids GE 6	0.191	(0.144)	0.271	(0.238)	0.273	(0.218)	0.635**	(0.305)	
Year 2002	-0.245**	(0.113)	-0.062	(0.212)	0.168	(0.199)	0.273	(0.290)	
Year 2003	-0.355***	(0.117)	0.183	(0.213)	-0.249	(0.199)	0.068	(0.299)	
Year 2004	-0.239*	(0.123)	0.574***	(0.218)	0.175	(0.215)	0.662**	(0.311)	
Employed in t=0	2.575***	(0.269)	0.108	(0.373)	1.313***	(0.424)	0.329	(0.490)	
Welfare receipt in t=0	0.028	(0.378)	2.232***	(0.486)	-0.429	(0.456)	2.007***	(0.648)	
M: Health status: good	0.509**	(0.206)	-0.711**	(0.357)	0.521	(0.348)	0.402	(0.517)	
M: No. of kids LT 6	-1.698***	(0.240)	-1.279**	(0.504)	-0.529	(0.373)	-0.312	(0.549)	
M: No. of kids GE 6	-0.128	(0.190)	-0.209	(0.311)	-0.367	(0.275)	-0.455	(0.385)	
Constant	-11.359***	(1.971)	-4.315	(3.561)	-5.808**	(2.826)	-2.375	(4.317)	
$\overline{Var(a_{ij})}$	2.803	(0.476)	1.647	(0.646)	1.245	(0.586)	1.771	(0.961)	
$Cov(a_{i,empl}, a_{i,welf})$	0.204 (0.509)				-0.276 (0.592)				
log likelihood	-3668.710				-1232.784				
No. of household-year observations	13,781				2,953				
No. of households		4,172	2		922				

*Note*: Dynamic multinomial logit models with random effects. Dependent variable: labor market state (inactivity, employment, welfare receipt). M: denotes individual-specific averages of a variable. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01.

Source: SOEP 2000-2004.

Table A3
Separate estimation results for natives and immigrants: post-reform

Variable		Nativo	es			Immigra	ants		
	Employme	ent	Welfare rec	eipt	Employme	ent	Welfare receipt		
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	
Employed in t-1	2.221***	(0.156)	0.523**	(0.264)	2.091***	(0.354)	0.695	(0.431)	
Welfare receipt in t-1	1.472***	(0.286)	1.941***	(0.326)	1.636***	(0.443)	1.877***	(0.456)	
Age	0.736***	(0.096)	0.830***	(0.184)	0.726***	(0.242)	-0.261	(0.295)	
Age squared	-0.875***	(0.106)	-0.934***	(0.203)	-0.913***	(0.266)	0.208	(0.324)	
Female	7.876***	(2.591)	18.760***	(4.910)	7.965	(5.953)	-3.037	(7.365)	
$Age \times Female$	-0.485***	(0.121)	-0.908***	(0.226)	-0.508*	(0.276)	0.092	(0.343)	
Age sq. × Female	0.600***	(0.137)	1.014***	(0.253)	0.629**	(0.308)	-0.092	(0.387)	
Education	0.047**	(0.023)	-0.177***	(0.046)	0.137**	(0.057)	-0.060	(0.079)	
School in Germany: no	_		_		0.101	(0.318)	-0.045	(0.416)	
Married	-0.685***	(0.136)	-1.906***	(0.237)	-0.675**	(0.332)	-1.051***	(0.394)	
Health status: good	-0.116	(0.131)	-0.911***	(0.226)	-0.359	(0.265)	-0.088	(0.337)	
No. of kids LT 6	0.388**	(0.161)	0.370	(0.291)	0.316	(0.356)	-0.095	(0.471)	
No. of kids GE 6	0.124	(0.146)	0.306	(0.248)	0.341	(0.308)	0.189	(0.373)	
Year 2007	0.180	(0.125)	-0.486**	(0.213)	0.054	(0.283)	-0.219	(0.352)	
Year 2008	0.330**	(0.133)	-0.332	(0.225)	0.469	(0.306)	-0.525	(0.394)	
Year 2009	0.226	(0.138)	-0.919***	(0.254)	0.276	(0.322)	-0.090	(0.400)	
Year 2010	0.313**	(0.146)	-0.199	(0.249)	0.061	(0.332)	0.175	(0.411)	
Employed in t=0	2.562***	(0.245)	-0.089	(0.374)	2.592***	(0.558)	-0.992	(0.607)	
Welfare receipt in t=0	0.371	(0.349)	3.203***	(0.507)	0.391	(0.523)	1.901***	(0.596)	
M: Health status: good	0.577**	(0.225)	-0.257	(0.403)	1.209**	(0.507)	-0.600	(0.647)	
M: No. of kids LT 6	-1.068***	(0.298)	0.110	(0.520)	-1.655***	(0.603)	-0.775	(0.719)	
M: No. of kids GE 6	0.265	(0.164)	-0.054	(0.294)	-0.246	(0.315)	0.461	(0.396)	
Constant	-15.202***	(2.086)	-16.502***	(4.077)	-14.694***	(5.345)	7.980	(6.519)	
$\overline{Var(a_{ij})}$	2.484	(0.401)	3.898	(0.963)	1.902	(0.897)	2.413	(1.144)	
$Cov(a_{i,empl}, a_{i,welf})$	0.022 (0.497)				-0.475 (0.717)				
log likelihood	-3456.030				-826.853				
No. of household-year observations	12,977				2,274				
No. of households	3,266								

*Note*: Dynamic multinomial logit models with random effects. Dependent variable: labor market state (inactivity, employment, welfare receipt). M: denotes individual-specific averages of a variable. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01.

Source: SOEP 2005-2010.

Table A4
Averages of predicted probabilities of labor market transitions given observed characteristics

State at time $t-1$				St	tate at time	et				
		Inactive		E	Employment			Welfare		
	Mean	95%	ó-CI	Mean	95%	5-CI	Mean	95%	ó-CI	
A. Natives: pre-reform	ı									
Inactive	0.237	0.183	0.304	0.724	0.655	0.777	0.039	0.022	0.070	
Employment	0.078	0.057	0.103	0.902	0.873	0.925	0.020	0.011	0.035	
Welfare	0.100	0.061	0.152	0.802	0.718	0.855	0.098	0.059	0.173	
B. Natives: post-reform	m									
Inactive	0.218	0.164	0.283	0.730	0.662	0.784	0.053	0.032	0.086	
Employment	0.067	0.047	0.091	0.902	0.871	0.927	0.032	0.019	0.050	
Welfare	0.091	0.058	0.137	0.823	0.755	0.869	0.086	0.055	0.139	
C. Immigrants: pre-ref	form									
Inactive	0.351	0.216	0.525	0.518	0.345	0.649	0.131	0.063	0.265	
Employment	0.097	0.051	0.163	0.860	0.778	0.916	0.043	0.018	0.093	
Welfare	0.134	0.061	0.246	0.665	0.457	0.780	0.201	0.108	0.408	
D. Immigrants: post-re	eform									
Inactive	0.225	0.119	0.394	0.663	0.498	0.769	0.112	0.051	0.219	
Employment	0.081	0.035	0.152	0.831	0.740	0.901	0.088	0.039	0.164	
Welfare	0.080	0.029	0.175	0.752	0.612	0.841	0.168	0.090	0.295	

*Note*: Calculations are based on estimation results in Tables A2 and A3. Simulation-based 95% confidence intervals are calculated using 1000 replications.

Table A5
Estimation results: regional unemployment rate (pre-reform)

Variable		Native	es		Immigrants				
	Employme	ent	Welfare rec	eipt	Employme	ent	Welfare rec	eipt	
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	
Employed in t-1	1.995***	(0.397)	1.174*	(0.689)	2.791***	(0.707)	0.797	(1.039)	
Welfare receipt in t-1	2.630***	(0.868)	4.047***	(0.787)	-0.076	(1.102)	1.381	(1.118)	
Employed in t-1 × unempl. Rate	0.025	(0.041)	-0.088	(0.067)	-0.030	(0.071)	-0.107	(0.104)	
Welfare receipt in t-1 $\times$ unempl. Rate	-0.125	(0.083)	-0.172**	(0.071)	0.172	(0.111)	0.033	(0.109)	
Unemployment rate	-0.063*	(0.037)	0.153***	(0.048)	-0.097	(0.062)	0.106	(0.073)	
Age	0.517***	(0.090)	0.339**	(0.163)	0.249*	(0.132)	-0.054	(0.198)	
Age squared	-0.642***	(0.100)	-0.403**	(0.180)	-0.337**	(0.148)	0.082	(0.218)	
Female	1.878	(2.459)	5.565	(4.312)	-4.375	(4.012)	0.217	(6.044)	
$Age \times Female$	-0.180	(0.115)	-0.269	(0.200)	0.106	(0.190)	-0.028	(0.284)	
Age sq. $\times$ Female	0.239*	(0.129)	0.279	(0.223)	-0.071	(0.218)	0.020	(0.322)	
Education	0.121***	(0.023)	-0.265***	(0.048)	0.113***	(0.039)	-0.000	(0.057)	
School in Germany: no	_		_		-0.240	(0.220)	0.280	(0.352)	
Married	-0.264**	(0.127)	-1.186***	(0.211)	-0.128	(0.240)	-0.463	(0.357)	
Health status: good	0.040	(0.123)	-0.313	(0.214)	-0.063	(0.209)	-0.805***	(0.303)	
No. of kids LT 6	0.690***	(0.156)	0.739**	(0.310)	0.134	(0.262)	0.678*	(0.372)	
No. of kids GE 6	0.196	(0.144)	0.267	(0.239)	0.296	(0.221)	0.644**	(0.305)	
Year 2002	-0.217*	(0.113)	-0.116	(0.214)	0.222	(0.201)	0.249	(0.291)	
Year 2003	-0.287**	(0.120)	0.052	(0.218)	-0.129	(0.205)	-0.019	(0.307)	
Year 2004	-0.163	(0.126)	0.445**	(0.222)	0.303	(0.222)	0.559*	(0.319)	
Employed in t=0	2.553***	(0.268)	0.141	(0.376)	1.324***	(0.419)	0.297	(0.492)	
Welfare receipt in t=0	0.067	(0.382)	2.153***	(0.480)	-0.373	(0.459)	1.951***	(0.645)	
M: Health status: good	0.496**	(0.205)	-0.730**	(0.359)	0.481	(0.350)	0.412	(0.518)	
M: No. of kids LT 6	-1.710***	(0.240)	-1.283**	(0.509)	-0.481	(0.376)	-0.344	(0.547)	
M: No. of kids GE 6	-0.149	(0.189)	-0.190	(0.313)	-0.382	(0.278)	-0.475	(0.386)	
Constant	-10.693***	(2.007)	-6.590*	(3.651)	-4.822*	(2.864)	-2.287	(4.334)	
$\overline{Var(a_{ij})}$	2.737	(0.268)	1.680	(0.693)	1.279	(0.770)	1.713	(0.746)	
$Cov(a_{i,empl}, a_{i,welf})$	0.278 (0.615)				-0.300 (0.574)				
log likelihood	-3656.795				-1225.318				
No. of household-year observations	13,781				2,953				
No. of households		4,172			922				

Note: Dynamic multinomial logit models with random effects. Robust standard errors clustered by region in parentheses. Dependent variable: labor market state (inactivity, employment, welfare receipt). M: denotes individual-specific averages of a variable. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01.

Source: SOEP 2000-2004.

Table A6
Estimation results: regional unemployment rate (post-reform)

Variable		Nativ	es		Immigrants			
	Employme	ent	Welfare rec	eipt	Employme	ent	Welfare receipt	
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.
Employed in t-1	2.268***	(0.373)	0.849	(0.648)	3.521***	(0.750)	2.464**	(1.028)
Welfare receipt in t-1	2.584***	(0.726)	3.202***	(0.781)	4.264***	(1.159)	3.391***	(1.157)
Employed in t-1 $\times$ unempl. Rate	-0.005	(0.038)	-0.035	(0.064)	-0.132*	(0.079)	-0.164	(0.111)
Welfare receipt in t-1 $\times$ unempl. Rate	-0.118*	(0.071)	-0.131*	(0.074)	-0.260**	(0.123)	-0.136	(0.113)
Unemployment rate	-0.006	(0.035)	0.092*	(0.053)	-0.042	(0.070)	0.138*	(0.080)
Age	0.733***	(0.096)	0.823***	(0.183)	0.714***	(0.234)	-0.418	(0.295)
Age squared	-0.871***	(0.106)	-0.926***	(0.201)	-0.896***	(0.259)	0.388	(0.322)
Female	7.810***	(2.592)	18.635***	(4.876)	8.417	(5.506)	-4.625	(7.047)
$Age \times Female$	-0.482***	(0.121)	-0.902***	(0.224)	-0.533**	(0.255)	0.179	(0.328)
Age sq. × Female	0.597***	(0.137)	1.008***	(0.251)	0.663**	(0.285)	-0.194	(0.369)
Education	0.049**	(0.023)	-0.180***	(0.046)	0.128**	(0.050)	-0.056	(0.069)
School in Germany: no	_		_		0.167	(0.292)	-0.214	(0.425)
Married	-0.680***	(0.136)	-1.864***	(0.235)	-0.513*	(0.296)	-0.836**	(0.360)
Health status: good	-0.120	(0.130)	-0.915***	(0.225)	-0.371	(0.253)	-0.059	(0.325)
No. of kids LT 6	0.393**	(0.161)	0.372	(0.290)	0.357	(0.340)	-0.015	(0.445)
No. of kids GE 6	0.127	(0.146)	0.308	(0.247)	0.393	(0.288)	0.275	(0.352)
Year 2007	0.154	(0.132)	-0.371*	(0.224)	-0.219	(0.284)	-0.154	(0.357)
Year 2008	0.286*	(0.150)	-0.155	(0.251)	-0.018	(0.316)	-0.426	(0.414)
Year 2009	0.188	(0.149)	-0.769***	(0.270)	-0.119	(0.323)	-0.001	(0.412)
Year 2010	0.272*	(0.160)	-0.029	(0.272)	-0.352	(0.343)	0.313	(0.436)
Employed in t=0	2.558***	(0.244)	-0.040	(0.375)	2.231***	(0.482)	-1.488***	(0.572)
Welfare receipt in t=0	0.408	(0.352)	3.145***	(0.504)	0.208	(0.427)	1.628***	(0.522)
M: Health status: good	0.591***	(0.225)	-0.218	(0.399)	1.108**	(0.472)	-0.641	(0.622)
M: No. of kids LT 6	-1.078***	(0.297)	0.111	(0.515)	-1.639***	(0.564)	-0.944	(0.675)
M: No. of kids GE 6	0.259	(0.164)	-0.035	(0.293)	-0.314	(0.287)	0.389	(0.374)
Constant	-15.100***	(2.106)	-17.313***	(4.064)	-14.136***	(5.059)	9.539	(6.459)
$\overline{Var(a_{ij})}$	2.476	(0.404)	3.643	(0.623)	1.197	(0.514)	1.784	(1.241)
$Cov(a_{i,empl}, a_{i,welf})$	-0.012		-1.462 (0.431)					
log likelihood	-3452.143				-819.835			
No. of household-year observations	12,977				2,274			
No. of households		3,26	6		616			

Note: Dynamic multinomial logit models with random effects. Robust standard errors clustered by region in parentheses. Dependent variable: labor market state (inactivity, employment, welfare receipt). M: denotes individual-specific averages of a variable. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01.

Source: SOEP 2000-2004.

Table A7
Predicted probabilities of labor market transitions
given subsample-specific average characteristics setting the initial condition to 2006

State at time $t-1$	State at time <i>t</i>									
	Inactive			E	Employmer	nt		Welfare		
	Mean	95%	o-CI	Mean	95%	o-CI	Mean	95%	ó-CI	
A. Natives, post-reform	1									
Inactive	0.184	0.143	0.244	0.790	0.727	0.829	0.026	0.016	0.050	
Employment	0.038	0.032	0.044	0.953	0.945	0.959	0.009	0.007	0.014	
Welfare	0.068	0.038	0.113	0.885	0.820	0.923	0.047	0.027	0.088	
B. Immigrants, post-re	form									
Inactive	0.149	0.085	0.281	0.779	0.633	0.855	0.073	0.039	0.156	
Employment	0.042	0.028	0.064	0.921	0.885	0.940	0.037	0.025	0.065	
Welfare	0.041	0.018	0.101	0.850	0.717	0.902	0.109	0.066	0.226	

*Note*: Simulation-based 95% confidence intervals are calculated using 1000 replications.

Table A8
Estimation results: alternative definition of states (total population, pre- and post-reform)

Variable		Pre-refo	orm		Post-reform			
	Employme	ent	Welfare rec	eipt	Employme	ent	Welfare rec	eipt
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.
Employed in t-1	2.329***	(0.140)	-0.218	(0.239)	2.196***	(0.139)	0.516**	(0.237)
Welfare receipt in t-1	1.310***	(0.257)	1.646***	(0.296)	0.991***	(0.258)	1.767***	(0.285)
Age	0.461***	(0.074)	0.046	(0.142)	0.724***	(0.087)	0.668***	(0.161)
Age squared	-0.579***	(0.082)	-0.037	(0.155)	-0.866***	(0.096)	-0.777***	(0.178)
Female	1.291	(2.044)	0.412	(3.910)	8.237***	(2.316)	15.247***	(4.253)
$Age \times Female$	-0.154	(0.096)	-0.008	(0.181)	-0.498***	(0.108)	-0.770***	(0.197)
Age sq. × Female	0.215**	(0.108)	-0.041	(0.203)	0.610***	(0.122)	0.873***	(0.221)
Education	0.105***	(0.019)	-0.210***	(0.043)	0.046**	(0.020)	-0.181***	(0.043)
School in Germany: no	-0.446***	(0.142)	0.784***	(0.231)	-0.021	(0.196)	0.832***	(0.301)
Married	-0.209*	(0.109)	-1.560***	(0.204)	-0.707***	(0.122)	-1.778***	(0.215)
Health status: good	-0.004	(0.105)	-0.453**	(0.194)	-0.183	(0.116)	-0.571***	(0.205)
No. of kids LT 6	0.598***	(0.132)	0.514*	(0.263)	0.395***	(0.145)	0.169	(0.263)
No. of kids GE 6	0.262**	(0.119)	0.417**	(0.209)	0.186	(0.130)	0.270	(0.215)
Year 2002	-0.156	(0.097)	0.189	(0.189)	_		_	
Year 2003	-0.291***	(0.100)	0.085	(0.197)	_		_	
Year 2004	-0.097	(0.105)	0.521***	(0.201)	_		_	
Year 2007	_		_		0.111	(0.113)	-0.186	(0.197)
Year 2008	_		_		0.304**	(0.121)	-0.252	(0.214)
Year 2009	_		_		0.152	(0.125)	-0.375	(0.229)
Year 2010	_		_		0.216	(0.132)	0.183	(0.228)
Employed in t=0	2.177***	(0.222)	0.095	(0.325)	2.386***	(0.215)	-0.586*	(0.356)
Welfare receipt in t=0	-0.119	(0.310)	2.611***	(0.422)	1.133***	(0.275)	2.375***	(0.354)
M: Health status: good	0.523***	(0.175)	-0.805**	(0.340)	0.612***	(0.202)	-0.217	(0.366)
M: No. of kids LT 6	-1.434***	(0.199)	-0.288	(0.407)	-1.133***	(0.261)	0.035	(0.448)
M: No. of kids GE 6	-0.277*	(0.155)	-0.105	(0.272)	0.181	(0.142)	0.040	(0.246)
Constant	-9.822***	(1.603)	-0.621	(3.149)	-14.574***	(1.892)	-12.341***	(3.546)
$\overline{Var(a_{ij})}$	2.243	(0.362)	2.063	(0.426)	2.247	(0.344)	3.438	(0.727)
$Cov(a_{i,empl}, a_{i,welf})$	-0.172		0.691 (0.461)					
log likelihood	-0.172 (0.592) -4624.9947				-4039.3653			
No. of household-year observations	16,734				15,251			
No. of households	5,094					3,882	2	

*Note*: Dynamic multinomial logit models with random effects. Dependent variable: labor market state (inactivity, employment, welfare receipt), *Aufstocker* are coded as employed. M: denotes individual-specific averages of a variable. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01. *Source*: SOEP 2000-2010.

Table A9
Estimation results: alternative definition of states (natives and immigrants, pre-reform)

Variable		Native	es		Immigrants			
	Employme	ent	Welfare rec	eipt	Employme	_	Welfare rec	eipt
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.
Employed in t-1	2.246***	(0.158)	0.196	(0.294)	2.747***	(0.295)	-0.853*	(0.443)
Welfare receipt in t-1	1.354***	(0.326)	1.852***	(0.371)	1.429***	(0.420)	1.216**	(0.504)
Age	0.530***	(0.089)	0.161	(0.191)	0.293**	(0.128)	-0.097	(0.230)
Age squared	-0.657***	(0.099)	-0.182	(0.209)	-0.389***	(0.145)	0.171	(0.252)
Female	2.830	(2.433)	0.519	(5.098)	-2.564	(3.833)	2.584	(6.951)
$Age \times Female$	-0.225**	(0.114)	-0.008	(0.236)	0.019	(0.182)	-0.099	(0.325)
Age sq. $\times$ Female	0.292**	(0.128)	-0.045	(0.264)	0.032	(0.209)	0.054	(0.367)
Education	0.110***	(0.022)	-0.338***	(0.064)	0.099***	(0.038)	0.003	(0.065)
School in Germany: no	_		_		-0.300	(0.213)	0.429	(0.415)
Married	-0.247**	(0.125)	-1.685***	(0.248)	-0.053	(0.231)	-0.926**	(0.411)
Health status: good	0.033	(0.122)	-0.355	(0.244)	-0.102	(0.206)	-0.742**	(0.337)
No. of kids LT 6	0.719***	(0.155)	0.536	(0.367)	0.249	(0.258)	0.743*	(0.414)
No. of kids GE 6	0.207	(0.143)	0.236	(0.280)	0.302	(0.215)	0.714**	(0.334)
Year 2002	-0.258**	(0.112)	0.163	(0.239)	0.169	(0.199)	0.192	(0.321)
Year 2003	-0.319***	(0.116)	0.128	(0.247)	-0.230	(0.198)	-0.081	(0.337)
Year 2004	-0.201*	(0.121)	0.565**	(0.251)	0.222	(0.213)	0.460	(0.351)
Employed in t=0 (2000)	2.474***	(0.263)	-0.366	(0.418)	1.046**	(0.409)	0.608	(0.536)
Welfare receipt in t=0 (2000)	0.057	(0.401)	2.677***	(0.524)	-0.804*	(0.479)	2.545***	(0.753)
M: Health status: good	0.519**	(0.204)	-1.043**	(0.427)	0.584*	(0.341)	-0.226	(0.597)
M: No. of kids LT 6	-1.720***	(0.237)	-0.827	(0.604)	-0.686*	(0.366)	0.070	(0.608)
M: No. of kids GE 6	-0.162	(0.187)	-0.051	(0.371)	-0.442	(0.272)	-0.401	(0.423)
Constant	-11.366***	(1.949)	-1.103	(4.268)	-6.290**	(2.758)	-1.492	(5.056)
$\overline{Var(a_{ij})}$	2.654	(0.454)	1.958	(0.705)	1.014	(0.547)	2.184	(1.130)
$Cov(a_{i,empl}, a_{i,welf})$	-0.263 (0.559)				-0.561 (0.645)			
log likelihood	-3458.4535				-1128.8688			
No. of household-year observations	13,781				2,953			
No. of households	4,172			922				

*Note*: Dynamic multinomial logit models with random effects. Dependent variable: labor market state (inactivity, employment, welfare receipt), *Aufstocker* are coded as employed. M: denotes individual-specific averages of a variable. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01. *Source*: SOEP 2000-2004.

Table A10
Estimation results: alternative definition of states (natives and immigrants, post-reform)

Variable		Nativ	es		Immigrants				
	Employme	ent	Welfare rec	eipt	Employme	_	Welfare rec	eipt	
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	
Employed in t-1	2.213***	(0.153)	0.706**	(0.293)	2.185***	(0.336)	0.304	(0.412)	
Welfare receipt in t-1	0.967***	(0.315)	1.884***	(0.346)	1.312***	(0.441)	1.761***	(0.469)	
Age	0.732***	(0.095)	1.006***	(0.219)	0.634***	(0.233)	-0.041	(0.281)	
Age squared	-0.868***	(0.105)	-1.126***	(0.240)	-0.805***	(0.255)	-0.037	(0.306)	
Female	8.240***	(2.550)	22.617***	(5.793)	6.844	(5.852)	-1.615	(7.097)	
$Age \times Female$	-0.495***	(0.119)	-1.101***	(0.264)	-0.447*	(0.270)	0.018	(0.329)	
Age sq. $\times$ Female	0.606***	(0.134)	1.239***	(0.293)	0.555*	(0.302)	-0.012	(0.369)	
Education	0.040*	(0.022)	-0.185***	(0.054)	0.126**	(0.055)	-0.061	(0.071)	
School in Germany: no	_		_		0.007	(0.306)	0.005	(0.371)	
Married	-0.734***	(0.133)	-2.166***	(0.276)	-0.772**	(0.320)	-1.112***	(0.364)	
Health status: good	-0.146	(0.130)	-0.934***	(0.260)	-0.346	(0.262)	-0.095	(0.338)	
No. of kids LT 6	0.428***	(0.159)	0.198	(0.329)	0.294	(0.353)	-0.016	(0.459)	
No. of kids GE 6	0.148	(0.144)	0.448	(0.281)	0.389	(0.304)	0.130	(0.364)	
Year 2007	0.122	(0.124)	-0.232	(0.245)	0.029	(0.280)	-0.141	(0.351)	
Year 2008	0.272**	(0.132)	-0.092	(0.257)	0.424	(0.302)	-0.510	(0.403)	
Year 2009	0.146	(0.137)	-0.532*	(0.283)	0.221	(0.317)	-0.024	(0.409)	
Year 2010	0.254*	(0.145)	0.102	(0.281)	0.047	(0.326)	0.249	(0.410)	
Employed in t=0	2.452***	(0.240)	-0.940*	(0.480)	2.177***	(0.503)	-0.257	(0.556)	
Welfare receipt in t=0	1.129***	(0.328)	2.625***	(0.452)	0.923*	(0.481)	1.698***	(0.512)	
M: Health status: good	0.510**	(0.222)	0.009	(0.460)	1.030**	(0.491)	-0.388	(0.605)	
M: No. of kids LT 6	-1.075***	(0.293)	0.404	(0.601)	-1.553***	(0.588)	-0.995	(0.694)	
M: No. of kids GE 6	0.264	(0.161)	-0.459	(0.344)	-0.232	(0.311)	0.394	(0.374)	
Constant	-14.878***	(2.060)	-20.454***	(4.893)	-12.269**	(5.177)	3.647	(6.260)	
$\overline{Var(a_{ij})}$	2.313	(0.382)	3.866	(0.960)	1.707	(0.794)	1.326	(0.823)	
$Cov(a_{i,empl}, a_{i,welf})$	0.158 (0.590)				0.555 (0.619)				
log likelihood	-3196.7524				-801.9247				
No. of household-year observations	12,977				2,274				
No. of households	3,266				616				

*Note*: Dynamic multinomial logit models with random effects. Dependent variable: labor market state (inactivity, employment, welfare receipt), *Aufstocker* are coded as employed. M: denotes individual-specific averages of a variable. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01.

Source: SOEP 2005-2010.

Table A11 Simulated predicted probabilities of labor market transitions for pre-reform characteristics and post-reform coefficients

State at time $t-1$	State at time <i>t</i>									
		Inactive		E	Employment			Welfare		
	Mean	Mean 95%-CI			95%	5-CI	Mean	95%	ó-CI	
A. Characteristics of to	otal population									
Inactive	0.206	0.172	0.251	0.740	0.694	0.777	0.054	0.040	0.073	
Employment	0.046	0.038	0.055	0.930	0.918	0.939	0.024	0.019	0.031	
Welfare	0.070	0.048	0.101	0.847	0.802	0.878	0.083	0.061	0.114	
B. Characteristics of na	natives									
Inactive	0.209	0.170	0.255	0.747	0.695	0.789	0.044	0.031	0.066	
Employment	0.046	0.038	0.054	0.936	0.925	0.946	0.018	0.014	0.025	
Welfare	0.072	0.046	0.111	0.851	0.792	0.887	0.078	0.051	0.121	
B. Characteristics of in	nmigrants									
Inactive	0.196	0.122	0.317	0.704	0.564	0.786	0.100	0.062	0.168	
Employment	0.044	0.027	0.072	0.902	0.853	0.929	0.054	0.034	0.092	
Welfare	0.052	0.023	0.108	0.802	0.688	0.861	0.146	0.095	0.244	

*Note*: Simulation-based 95% confidence intervals are calculated using 1000 replications. Calculations are based on estimation results for post-reform period in Tables 4 and A3.

Table A12
Estimation results: including initial period explanatory variables

Variable		Pre refe	om		Post-reform				
	Employme	ent	Welfare rec	eipt	Employme	ent	Welfare receipt		
	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	Coef.	S.E.	
Employed in t-1	2.267***	(0.143)	0.065	(0.219)	2.182***	(0.140)	0.554**	(0.223)	
Welfare receipt in t-1	1.413***	(0.242)	2.048***	(0.281)	1.481***	(0.240)	1.819***	(0.270)	
Age	0.452***	(0.075)	0.173	(0.125)	0.728***	(0.088)	0.535***	(0.151)	
Age squared	-0.568***	(0.084)	-0.206	(0.137)	-0.874***	(0.098)	-0.630***	(0.167)	
Female	0.503	(2.080)	2.939	(3.407)	7.915***	(2.357)	13.449***	(3.980)	
$Age \times Female$	-0.118	(0.097)	-0.151	(0.158)	-0.488***	(0.110)	-0.667***	(0.185)	
Age sq. × Female	0.172	(0.110)	0.150	(0.178)	0.604***	(0.124)	0.747***	(0.208)	
Education	0.116***	(0.019)	-0.179***	(0.036)	0.059***	(0.021)	-0.177***	(0.040)	
School in Germany: no	-0.416***	(0.145)	0.577***	(0.206)	-0.029	(0.201)	0.649**	(0.290)	
Married	-0.212*	(0.111)	-1.109***	(0.177)	-0.619***	(0.124)	-1.682***	(0.206)	
Health status: good	0.027	(0.106)	-0.466***	(0.173)	-0.139	(0.118)	-0.638***	(0.188)	
No. of kids LT 6	0.539***	(0.134)	0.602**	(0.236)	0.395***	(0.146)	0.206	(0.246)	
No. of kids GE 6	0.267**	(0.121)	0.406**	(0.188)	0.151	(0.132)	0.201	(0.203)	
Year 2002	-0.141	(0.098)	0.056	(0.170)	_	` ,	_	` ,	
Year 2003	-0.317***	(0.101)	0.146	(0.173)	_		_		
Year 2004	-0.134	(0.106)	0.572***	(0.178)	_		_		
Year 2007	_		_	, ,	0.161	(0.115)	-0.377**	(0.181)	
Year 2008	_		_		0.351***	(0.122)	-0.402**	(0.195)	
Year 2009	_		_		0.229*	(0.127)	-0.673***	(0.212)	
Year 2010	_				0.276**	(0.133)	-0.056	(0.211)	
Employed in t=0	2.337***	(0.230)	0.252	(0.301)	2.554***	(0.219)	-0.369	(0.312)	
Welfare receipt in t=0	0.027	(0.295)	2.298***	(0.396)	0.383	(0.292)	2.896***	(0.407)	
M: Health status: good	0.517**	(0.215)	-0.443	(0.353)	0.482*	(0.262)	0.015	(0.440)	
M: No. of kids LT 6	-1.499***	(0.209)	-0.813**	(0.376)	-1.700***	(0.329)	-0.163	(0.524)	
M: No. of kids GE 6	-0.348**	(0.172)	-0.157	(0.266)	0.368	(0.260)	-0.111	(0.410)	
I: Health status: good	-0.028	(0.126)	0.044	(0.197)	0.186	(0.152)	-0.424	(0.258)	
I: No. of kids LT 6	0.305**	(0.126)	0.061	(0.189)	0.441**	(0.179)	0.199	(0.294)	
I: No. of kids GE 6	0.023	(0.087)	-0.172	(0.127)	-0.126	(0.188)	0.290	(0.298)	
Constant	-9.806***	(1.630)	-3.030	(2.740)	-14.992***	(1.922)	-9.510***	(3.316)	
$\overline{Var(a_{ij})}$	2.409	(0.384)	1.879	(0.559)	2.396	(0.359)	3.991	(0.824)	
$Cov(a_{i,empl}, a_{i,welf})$	0.094 (0.398)				0.054 (0.408)				
log likelihood	-4927.576			-4297.9657					
No. of household-year observations	16,718			15,215					
No. of households		5,077			3,860				

*Note*: Dynamic multinomial logit models with random effects. Dependent variable: labor market state (inactivity, employment, welfare receipt). M: denotes individual-specific averages of a variable. I: denotes initial-period explanatory variable. The numbers of individuals differ from those in Table 4 because of missing values of initial-period explanatory variables. Significance level: \*<0.1, \*\*<0.05, \*\*\*<0.01.

Source: SOEP 2000-2010.

Table A13
Predicted probabilities of labor market transitions given subsample-specific average characteristics (estimations including initial period explanatory variables)

State at time $t-1$		State at time <i>t</i>									
		Inactive		E	mployme	nt		Welfare			
	Mean	95%	5-CI	Mean	95%	5-CI	Mean	95%	5-CI		
A. Total population, pre-reform											
Inactive	0.244	0.205	0.299	0.719	0.663	0.759	0.037	0.026	0.054		
Employment	0.055	0.049	0.061	0.936	0.929	0.942	0.009	0.008	0.012		
Welfare	0.089	0.062	0.121	0.828	0.770	0.869	0.083	0.057	0.129		
B. Total population, po	st-reform										
Inactive	0.195	0.164	0.239	0.767	0.724	0.799	0.038	0.028	0.052		
Employment	0.042	0.037	0.047	0.942	0.936	0.949	0.016	0.013	0.020		
Welfare	0.066	0.046	0.093	0.873	0.835	0.900	0.061	0.045	0.086		

*Note*: Simulation-based 95% confidence intervals are calculated using 1000 replications. Calculations are based on estimation results in Table A12.