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A uniform central limit theorem and efficiency for deconvolution estimators

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Abstract

We estimate linear functionals in the classical deconvolution problem by kernel estimators. We obtain a uniform central limit theorem with \sqrt{n} -rate on the assumption that the smoothness of the functionals is larger than the ill-posedness of the problem, which is given by the polynomial decay rate of the characteristic function of the error. The limit distribution is a generalized Brownian bridge with a covariance structure that depends on the characteristic function of the error and on the functionals. The proposed estimators are optimal in the sense of semiparametric efficiency. The class of linear functionals is wide enough to incorporate the estimation of distribution functions. The proofs are based on smoothed empirical processes and mapping properties of the deconvolution operator.

Keywords: Deconvolution · Donsker theorem · Efficiency · Distribution function · Smoothed empirical processes · Fourier multiplier

MSC (2000): 62G05 · 60F05

JEL Classification: C14

1 Introduction

Our observations are given by $n \in \mathbb{N}$ independent and identically distributed random variables

$$Y_j = X_j + \varepsilon_j, \quad j = 1, \dots, n, \quad (1)$$

where X_j and ε_j are independent of each other, the distribution of the errors ε_j is supposed to be known and the aim is statistical inference on the distribution of X_j . Let us denote the densities of X_j and ε_j by f_X and f_ε , respectively. We consider the case of ordinary smooth errors, which means that the characteristic function φ_ε of the errors ε_j decays with polynomial

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rate, determining the ill-posedness of the inverse problem. The contribution of this article to the well studied problem of deconvolution is twofold. First, we prove a uniform central limit theorem for kernel estimators of the distribution function of X_j in the setting of \sqrt{n} convergence rates. More precisely, the theorem does not only include the estimation of the distribution function but covers translation classes of linear functionals of the density f_X whenever the ill-posedness is smaller than the smoothness of the functionals. Second, we obtain more exact results than the minimax rates of convergence by showing that the used estimators are optimal in the sense of semiparametric efficiency.

The classical Donsker theorem plays a central role in statistics and states that the empirical distribution function of an independent, identically distributed sample converges uniformly to the distribution function. In the deconvolution model (1) our Donsker theorem states uniform convergence for an asymptotically unbiased estimator of translated functionals $t \mapsto \vartheta_t := \int \zeta(x - t) f_X(x) dx$, where the special case $\zeta := \mathbb{1}_{(-\infty, 0]}$ leads to the estimation of the distribution function. This generalization allows to consider functionals ϑ_t as long as the smoothness of ζ in an L^2 -Sobolev sense compensates the ill-posedness of the problem. The limiting process \mathbb{G} in the uniform central limit theorem is a generalized Brownian bridge, whose covariance depends on the functional ζ and through the deconvolution operator $\mathcal{F}^{-1}[1/\varphi_\varepsilon]$ also on the distribution of the errors. The used kernel estimators $\hat{\vartheta}_t$ are minimax optimal since they converge with a \sqrt{n} -rate. So investigating optimality further leads naturally to the question whether the asymptotic variance of the estimators is minimal, as in the case of the empirical distribution function in the classical Donsker theorem. We prove that the estimator $\hat{\vartheta}_\bullet$ is efficient in the sense of a Hájek-Le Cam convolution theorem. In particular, the asymptotic covariance matrices of the finite dimensional distributions achieve the Cramér-Rao information bound. By uniform convergence and efficiency the kernel estimator of f_X fulfills the ‘plug-in’ property of Bickel and Ritov [2] in the deconvolution model (1).

The deconvolution problem has attracted much attention so we mention here only closely related works and refer the interested reader to the references therein. The classical works by Fan [11, 12] contain asymptotic normality of kernel density estimators as well as minimax convergence rates for estimating the density and the distribution function. Butucea and Comte [5] have treated the data-driven choice of the bandwidth for estimating functionals of f_X but assumed some minimal smoothness and integrability conditions on the functional ϑ_t , which exclude, for example, $\zeta := \mathbb{1}_{(-\infty, 0]}$ since it is not integrable. Dattner et al. [6] have studied minimax-optimal and adaptive estimation of the distribution function. Asymptotic normality of estimators for the distribution function has been shown by van Es and Uh [31] in the case of supersmooth errors and by Hall and Lahiri [18] for ordinary smooth errors. In contrast we consider the estimation of general linear functionals and are interested in uniform convergence. Uniform results have been studied for the density but not for the distribution function by Bissantz et al. [3] and by Lounici and Nickl [21]. Recently, Nickl and Reiß [24] have proved a Donsker theorem for estimators of the distribution function of a Lévy measure. Their situation is related but more involved than ours, owing to the nonlinearity and the auto-deconvolution of the Lévy measure. In a deconvolution context we consider the more general problem of estimating linear functionals efficiently, which contains estimating of the distribution function as a special case and provides clear insight in the interplay between smoothness of ζ and the ill-posedness of the problem. While efficiency has been investigated in various semiparametric models, e.g., see Bickel et al. [1], to the best of the authors knowledge there are no results in this direction in the deconvolution framework. However, in the Lévy setting Nickl and Reiß [24] have shown heuristically that their estimator achieves the lower bound of the variance

while a rigorous proof remained open.

In order to show the uniform central limit theorem in the deconvolution problem, we prove that the empirical process $\sqrt{n}(\mathbb{P}_n - \mathbb{P})$ is tight in the space of bounded functions acting on the class

$$\mathcal{G} := \{\mathcal{F}^{-1}[1/\varphi_\varepsilon(-\bullet)] * \zeta_t \mid t \in \mathbb{R}\}, \quad \zeta_t := \zeta(\bullet - t),$$

where \mathbb{P} and $\mathbb{P}_n = \frac{1}{n} \sum_{j=1}^n \delta_{Y_j}$ denote the true and the empirical probability measure of the observations Y_j , respectively. Since \mathcal{G} may consist of translates of an unbounded function, this is in general not a Donsker class. Nevertheless, Radulović and Wegkamp [26] have observed that a smoothed empirical processes might converge even when the unsmoothed process does not. Giné and Nickl [14] have further developed these ideas and have shown uniform central limit theorems for kernel density estimators. Nickl and Reiß [24] used smoothed empirical processes in the inverse problem of estimating the distribution function of Lévy measures. In order to show semiparametric efficiency in the deconvolution problem, the main problem is to show that the efficient influence function is indeed an element of the tangent space. If the regularity of ζ is small, the standard methods given in the monograph of Bickel et al. [1] do not apply in this ill-posed problem. Instead, we approximate ζ by a sequence of smooth (ζ_n) and show the convergence of the information bounds. Interestingly, this reveals a relation between the intrinsic metric of the limit \mathbb{G} and the metric which is induced by the inverse Fisher information. Additionally to techniques of smoothed empirical processes and the calculus of information bounds, our proofs rely on the Fourier multiplier property of the underlying deconvolution operator $\mathcal{F}^{-1}[1/\varphi_\varepsilon]$, which is related to pseudo-differential operators as noted in the Lévy process setting by Nickl and Reiß [24] and in the deconvolution context by Schmidt–Hieber et al. [27]. Important for our proofs are the mapping properties of $\mathcal{F}^{-1}[1/\varphi_\varepsilon]$ on Besov spaces.

This paper is organized as follows: In Section 2 we formulate the Donsker theorem and discuss its consequences. Efficiency is then considered in Section 3. All proofs are deferred to Sections 4 and 5. In the Appendix we summarize definitions and properties of the function spaces used in the paper.

2 Uniform central limit theorem

2.1 The estimator

According to the observation scheme (1), Y_j are distributed with density $f_Y = f_X * f_\varepsilon$ determining the probability measure \mathbb{P} . The characteristic function φ of \mathbb{P} can be estimated by its empirical version $\varphi_n(u) = \frac{1}{n} \sum_{j=1}^n e^{iuY_j}$, $u \in \mathbb{R}$. For ζ to be specified later and recalling $\zeta_t = \zeta(\bullet - t)$, our aim is to estimate functionals of the form

$$\vartheta_t := \langle \zeta_t, f_X \rangle = \int \zeta_t(x) f_X(x) dx. \quad (2)$$

Defining the Fourier transform by $\mathcal{F}f(u) := \int e^{iux} f(x) dx$, $u \in \mathbb{R}$, the natural estimator of the functional ϑ_t is given by

$$\hat{\vartheta}_t := \int \zeta_t(x) \mathcal{F}^{-1} \left[\mathcal{F} K_h \frac{\varphi_n}{\varphi_\varepsilon} \right] (x) dx, \quad (3)$$

where K is a kernel, $h > 0$ the bandwidth and we have written as usual $K_h(x) = h^{-1}K(x/h)$. Choosing $\mathcal{F}K = \mathbb{1}_{[-\pi, \pi]}$ for some $\pi > 0$ leads to the estimator proposed by Butucea and Comte [5]. Throughout, we suppose that

- (i) $K \in L^1(\mathbb{R}) \cap L^\infty(\mathbb{R})$ is symmetric and band-limited with $\text{supp}(\mathcal{F}K) \subseteq [-1, 1]$,
- (ii) for $l = 1, \dots, L$

$$\int K = 1, \quad \int x^l K(x) dx = 0, \quad \int |x^{L+1} K(x)| dx < \infty \quad \text{and} \quad (4)$$

- (iii) $K \in C^1(\mathbb{R})$ satisfies, denoting $\langle x \rangle := (1 + x^2)^{1/2}$,

$$|K(x)| + |K'(x)| \lesssim \langle x \rangle^{-2}. \quad (5)$$

Throughout, we write $A_p \lesssim B_p$ if there exists a constant $C > 0$ independent of the parameter p such that $A_p \leq CB_p$. If $A_p \lesssim B_p$ and $B_p \lesssim A_p$, we write $A_p \sim B_p$. Examples of such kernels can be obtained by taking $\mathcal{F}K$ to be a symmetric function in $C^\infty(\mathbb{R})$ which is supported in $[-1, 1]$ and constant to one in a neighborhood of zero. The resulting kernels are called flat top kernels and were used in deconvolution problems, for example, by Bissantz et al. [3].

2.2 Statement of the theorem

Given a function ζ specified later, our aim is to show a Donsker theorem for the estimator over the class of translations ζ_t , $t \in \mathbb{R}$. In view of the classical Donsker theorem in a model without additive errors, where no assumptions on the smoothness of the distribution are needed, we want to assume as less smoothness of f_X as possible still guaranteeing \sqrt{n} -rates. For some $\delta > 0$ the following assumptions on the density f_X will be needed:

Assumption 1.

- (i) Let f_X be bounded and assume the moment condition $\int |x|^{2+\delta} f_X(x) dx < \infty$.
- (ii) Assume $f_X \in H^\alpha(\mathbb{R})$ that is the density has Sobolev smoothness of order $\alpha \geq 0$.

We refer to the appendix for an exact definition of the Sobolev space $H^\alpha(\mathbb{R})$. Boundedness of the observation density f_Y follows immediately from (i) since $\|f_Y\|_\infty \leq \|f_X\|_\infty \|f_\varepsilon\|_{L^1} < \infty$. In addition to the smoothness of f_X , the smoothness of ζ will be crucial. We assume for $\gamma_s, \gamma_c > 0$

$$\begin{aligned} \zeta \in Z^{\gamma_s, \gamma_c} := \left\{ \zeta = \zeta^c + \zeta^s \mid \zeta^s \in H^{\gamma_s}(\mathbb{R}) \text{ is compactly supported as well} \right. \\ \left. \text{as } \langle x \rangle^\tau (\zeta^c(x) - a(x)) \in H^{\gamma_c}(\mathbb{R}) \text{ for some } \tau > 0 \text{ and} \right. \\ \left. \text{some } a \in C^\infty(\mathbb{R}) \text{ such that } a' \text{ is compactly supported} \right\} \end{aligned} \quad (6)$$

and write for $\zeta \in Z^{\gamma_s, \gamma_c}$ with a given decomposition $\zeta = \zeta^s + \zeta^c$

$$\|\zeta\|_{Z^{\gamma_s, \gamma_c}} := \|\zeta^s\|_{H^{\gamma_s}} + \left\| \frac{1}{ix+1} \zeta^c(x) \right\|_{H^{\gamma_c}},$$

which is finite since $\left\| \frac{1}{ix+1} \zeta^c(x) \right\|_{H^{\gamma_c}}$ is bounded by $\left\| \frac{a(x)}{ix+1} \right\|_{H^{\gamma_c}} + \left\| \frac{1}{(ix+1)\langle x \rangle^\tau} \right\|_{C^s} \|\langle x \rangle^\tau (\zeta^c(x) - a(x))\|_{H^{\gamma_c}} < \infty$ for any $s > \gamma_c$. Several examples for ζ and corresponding γ_s, γ_c will be given in Examples 1-3 below. In particular, $\mathbb{1}_{(-\infty, 0]} \in Z^{\gamma_s, \gamma_c}$ for $\gamma_s < 1/2$. The ill-posedness of the problem is determined by the decay of the characteristic function of the errors. More precisely, we suppose

Assumption 2. *Let the error distribution satisfy*

- (i) $\int |x|^{2+\delta} f_\varepsilon(x) dx < \infty$ thus φ_ε is twice continuously differentiable and
- (ii) $|(\varphi_\varepsilon^{-1})'(u)| \lesssim \langle u \rangle^{\beta-1}$ for some $\beta > 0$, in particular $|\varphi_\varepsilon^{-1}(u)| \lesssim \langle u \rangle^\beta, u \in \mathbb{R}$.

Throughout, we write $\varphi_\varepsilon^{-1} = 1/\varphi_\varepsilon$. The Assumption (ii) on the distribution of the errors is similar to the classical decay assumption by Fan [11] and it is fulfilled for many ordinary smooth error laws such as gamma or Laplace distributions as discussed below. Assumption 2(ii) implies that φ_ε^{-1} is a Fourier multiplier on Besov spaces so that

$$B_{p,q}^s(\mathbb{R}) \ni f \mapsto \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F} f] \in B_{p,q}^{s-\beta}(\mathbb{R})$$

for $p, q \in [1, \infty], s \in \mathbb{R}$, is a continuous linear map, which is essential in our proofs, compare Lemma 5. In the same spirit Schmidt–Hieber et al. [27] discuss the behavior of the deconvolution operator as pseudo-differential operator. We define

$$g_t := \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta_t \quad \text{and} \quad \mathcal{G} = \{g_t | t \in \mathbb{R}\}. \quad (7)$$

Note that in general g_t may only exist in a distributional sense, but on Assumption 2 and for $\zeta \in Z^{\gamma_s, \gamma_c}$ it can be rigorously interpreted by (see (19))

$$\begin{aligned} g_0(x) &= \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta^s(u)](x) \\ &\quad + (1 + ix) \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F}[\frac{1}{iy+1} \zeta^c(y)](u)](x) \\ &\quad + \mathcal{F}^{-1}[(\varphi_\varepsilon^{-1})'(-u) \mathcal{F}[\frac{1}{iy+1} \zeta^c(y)](u)](x), \end{aligned}$$

which indicates why we have imposed an assumption on $(\varphi_\varepsilon^{-1})'$ and have defined $\|\bullet\|_{Z^{\gamma_s, \gamma_c}}$ as above.

It will turn out that \mathcal{G} is \mathbb{P} -pregaussian, but not Donsker in general. Denoting by $\lfloor \alpha \rfloor$ the largest integer smaller or equal to α and defining convergence in law on $\ell^\infty(\mathbb{R})$ as Dudley [9, p. 94], we state our main result

Theorem 1. *Grant Assumptions 1 and 2 as well as $\zeta \in Z^{\gamma_s, \gamma_c}$ with $\gamma_s > \beta, \gamma_c > (1/2 \vee \alpha) + \gamma_s$ and $\alpha + 3\gamma_s > 2\beta + 1$. Furthermore, let the kernel K satisfy (4) with $L = \lfloor \alpha + \gamma_s \rfloor$. Let $h_n^{2\alpha+2\gamma_s} n \rightarrow 0$ and if $\gamma_s \leq \beta + 1/2$ let in addition $h_n^\rho n \rightarrow \infty$ for some $\rho > 4\beta - 4\gamma_s + 2$, then*

$$\sqrt{n}(\hat{\vartheta}_t - \vartheta_t)_{t \in \mathbb{R}} \xrightarrow{\mathcal{L}} \mathbb{G} \quad \text{in } \ell^\infty(\mathbb{R})$$

as $n \rightarrow \infty$, where \mathbb{G} is a centered Gaussian Borel random variable in $\ell^\infty(\mathbb{R})$ with covariance function given by

$$\Sigma_{s,t} := \int g_s(x) g_t(x) \mathbb{P}(dx) - \vartheta_s \vartheta_t$$

for g_s, g_t defined in (7) and $s, t \in \mathbb{R}$.

We illustrate the range of this theorem by the following examples.

Example 1. We consider the indicator function $\mathbb{1}_{(-\infty, 0]}(x)$, $x \in \mathbb{R}$. Let a be a monotone decreasing $C^\infty(\mathbb{R})$ function, which is for some $M > 0$ equal to zero for all $x \geq M$ and equal to one for all $x \leq -M$. We define $\zeta^s := \mathbb{1}_{(-\infty, 0]} - a$ and $\zeta^c := a$. From the bounded variation of ζ^s follows $\zeta^s \in B_{1,\infty}^1(\mathbb{R}) \subseteq H^{\gamma_s}(\mathbb{R})$ for any $\gamma_s < 1/2$ by Besov smoothness of

bounded variation functions (51) as well as by the Besov space embeddings (46) and (47). Since $a \in C^\infty(\mathbb{R})$ and a' is compactly supported, the condition on ζ^c is satisfied for any $\gamma_c > 0$. Hence, $\mathbb{1}_{(-\infty, t]} \in Z^{\gamma_s, \gamma_c}$ if $\gamma_s < 1/2$. On the other hand, this cannot hold for $\gamma_s > 1/2$ since $H^{\gamma_s}(\mathbb{R}) \subseteq C^0(\mathbb{R})$ by Sobolev's embedding theorem or by (45), (46) and (47). Owing to the condition $\gamma_s > \beta$, Assumption 2 needs to be fulfilled for some $\beta < 1/2$ which is done, for example, by the gamma distribution $\Gamma(\beta, \eta)$ with $\beta \in (0, 1/2)$ and $\eta \in (0, \infty)$, that is

$$f_\varepsilon(x) := \gamma_{\beta, \eta}(x) := \frac{1}{\Gamma(\beta)\eta^\beta} x^{\beta-1} e^{-x/\eta} \mathbb{1}_{[0, \infty)}(x), \quad x \in \mathbb{R},$$

and $\varphi_\varepsilon(u) = (1 - i\eta u)^{-\beta}$, $u \in \mathbb{R}$.

Example 2. Let $\zeta_t(x) := \zeta_t^s(x) := \max(K - |x - t|, 0)$ and $\zeta_t^c(x) := 0$ with $K > 0$. The payoff of the butterfly spread is described by such a function [13]. Then $\mathcal{F}\zeta(u) = 4 \sin^2(u/2)/u^2$ and $\zeta^s \in H^{\gamma_s}(\mathbb{R})$ for any $\gamma_s < 3/2$. So, Assumption 2 is required for some $\beta < 3/2$, which holds, for example, for the chi-squared distribution with one or two degrees of freedom or for the exponential distribution.

Example 3. Butucea and Comte [5] studied the case $\beta > 1$ and derived \sqrt{n} -rates for $\gamma_s > \beta$ in our notation. In particular, they considered supersmooth ζ , that is $\mathcal{F}\zeta$ decays exponentially. In this case $\zeta \in H^s(\mathbb{R})$ for any $s \in \mathbb{N}$. Requiring the slightly stronger assumption that $\langle x \rangle^\tau \zeta(x) \in H^s(\mathbb{R})$ for some arbitrary small $\tau > 0$ and for all $s \in \mathbb{N}$ we can choose $\zeta^c := \zeta$ and $\zeta^s := 0$. Then β can be taken arbitrary large such that all gamma distributions, the Laplace distributions and convolutions of them can be chosen as error distributions.

2.3 Discussion

To have \sqrt{n} -rates we suppose $\gamma_s > \beta$ in Theorem 1, which means that the smoothness of the functionals compensates the ill-posedness of the problem. This condition is natural in view of the abstract analysis in terms of Hilbert scales by Goldenshluger and Pereverzev [17], who obtain the minimax rate $n^{-(\alpha+\gamma_s)/(2\alpha+2\beta)} \vee n^{-1/2}$ in our notation. As a consequence of the condition on γ_s and γ_c we can bound the stochastic error term of the estimator $\hat{\vartheta}_t$ uniformly in $h \in (0, 1)$. The bias term is of order $h^{\alpha+\gamma_s}$.

For $\gamma_s > \beta + 1/2$ the class \mathcal{G} is a Donsker class. In this case the only condition on the bandwidth is that the bias tends faster than $n^{-1/2}$ to zero. In the interesting but involved case $\gamma_s \in (\beta, \beta + 1/2]$, the class \mathcal{G} will in general not be a Donsker class. Estimating the distribution function as in Example 1 belongs to this case. In order to see that \mathcal{G} is in general not a Donsker class, let the error distribution be given by $f_\varepsilon = \gamma_{\beta, \eta}(-\bullet)$ and $\zeta = \gamma_{\sigma, \eta}$ with $\sigma \in (\gamma_s + 1/2, \beta + 1)$. Then g_t equals $\gamma_{\sigma-\beta, \eta} * \delta_t$. For the shape parameter holds $\sigma - \beta \in (1/2, 1)$ and thus g_t is an $L^2(\mathbb{R})$ -function unbounded at t . The Lebesgue density of \mathbb{P} is bounded by Assumption 1(i). Hence, \mathcal{G} consists of all translates of an unbounded function and thus cannot be Donsker, cf. Theorem 7 by Nickl [22].

Therefore, for $\gamma_s \in (\beta, \beta + 1/2]$ smoothed empirical processes are necessary, especially we need to ensure enough smoothing to be able to obtain a uniform central limit theorem. The bandwidth cannot tend too fast to zero, more precisely we require $h_n^\rho n \rightarrow \infty$ as $n \rightarrow \infty$ for some ρ with $\rho > 4\beta - 4\gamma_s + 2$. In combination with the bias condition $h_n^{2\alpha+2\gamma_s} n \rightarrow 0$ as $n \rightarrow \infty$ we obtain necessarily $\alpha + \gamma_s > 2\beta - 2\gamma_s + 1$ leading to the assumption in the theorem. Since $2\alpha + 2\gamma_s > \alpha + 2\beta - \gamma_s + 1 > 4\beta - 4\gamma_s + 2$ we can always choose $h_n \sim n^{-1/(\alpha+2\beta-\gamma_s+1)}$. In contrast to Butucea and Comte [5], Dattner et al. [6], Fan [12] our

choice of the bandwidth h_n is not determined by the bias–variance trade–off, but rather by the amount of smoothing necessary to obtain a uniform central limit theorem. The classical bandwidth $h_n \sim n^{-1/(2\alpha+2\beta)}$ is optimal for estimating the density in the sense that it achieves the minimax rate with respect to the mean integrated squared error (MISE), compare Fan [12] who assumes Hölder smoothness of f_X instead of L^2 –Sobolev smoothness. For this choice the bias condition $h_n^{2\alpha+2\gamma_s} n \rightarrow 0$ is satisfied. If $\gamma_s \leq \beta + 1/2$ the classical bandwidth satisfies the additional minimal smoothness condition in the case of estimating the distribution function with mild conditions on f_X . It suffices for example that f_X is of bounded variation. Then α and γ_s can be chosen large enough in $(0, 1/2)$ such that $2\alpha+2\beta > 4\beta-4\gamma_s+2$ and the classical bandwidth satisfies the conditions of the theorem. Whenever the classical bandwidth $h_n \sim n^{-1/(2\alpha+2\beta)}$ satisfies the conditions of Theorem 1, then the corresponding density estimator is a ‘plug–in’ estimator in the sense of Bickel and Ritov [2] meaning that the density is estimated rate optimal for the MISE, the functionals are estimated efficiently (see Section 3) and the estimators of the functionals converge uniformly over $t \in \mathbb{R}$.

The smoothness condition on the density f_X is then a consequence of the given choice of h_n together with the classical bias estimate for kernel estimators. As we have seen in Example 1 for estimating the distribution function we have $\zeta = \mathbb{1}_{(-\infty, 0]} \in Z^{\gamma_s, \gamma_c}$ with $\gamma_s < 1/2$ arbitrary close to $1/2$. In the classical Donsker theorem which corresponds to the case $\beta \rightarrow 0$ the condition $\alpha + 3\gamma_s > 2\beta + 1$ would simplify to $\alpha > -1/2$. However, we suppose f_X to be bounded, which leads to much clearer proofs, and thus $f_X \in H^0(\mathbb{R})$ is automatically satisfied. Assumption 1 allows to focus on the interplay between the functional ζ and the deconvolution operator $\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}]$. Nickl and Reiß [24] have studied the case of unbounded densities, which is necessary in the Lévy process setup, but considered $\zeta_t = \mathbb{1}_{(-\infty, t]}$ only. The class Z^{γ_s, γ_c} is defined by L^2 –Sobolev conditions so that bounded variation arguments for ζ have to be avoided in the proofs.

An interesting aspect is the following: If we restrict the uniform convergence to $(\zeta_t)_{t \in T}$ for some compact set $T \subseteq \mathbb{R}$, it is sufficient to assume $\frac{1}{ix+1}\zeta^c \in H^{\gamma_c}(\mathbb{R})$ instead of requiring $(1 \vee |x|^\tau)(\zeta^c(x) - a(x)) \in H^{\gamma_c}(\mathbb{R})$ for some $\tau > 0$ and a function $a \in C^\infty(\mathbb{R})$ such that a' is compactly supported as done in Z^{γ_s, γ_c} . In particular, slowly growing ζ would be allowed. The stronger condition in the definition of Z^{γ_s, γ_c} is only needed to ensure polynomial covering numbers of $\{g_t | t \in T\}$ for $T \subseteq \mathbb{R}$ unbounded (cf. Theorem 7 below).

As a corollary of Theorem 1 we can weaken Assumption 2(ii). If the characteristic function of the errors ε is given by $\tilde{\varphi}_\varepsilon = \varphi_\varepsilon \psi$ where φ_ε satisfies Assumption 2(ii) and there is a Schwartz distribution $\nu \in \mathcal{S}'(\mathbb{R})$ such that $\mathcal{F}\nu = \psi^{-1}$ and $\nu * \zeta \in Z^{\gamma_s, \gamma_c}$ for $\zeta \in Z^{\gamma_s, \gamma_c}$, then for $t \in \mathbb{R}$

$$\mathcal{F}^{-1}[\tilde{\varphi}_\varepsilon^{-1}] * \zeta(\bullet - t) = \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * (\nu * \zeta)(\bullet - t)$$

and thus we can proceed as before. For instance, for translated errors $f_\varepsilon * \delta_\mu$ with $\mu \neq 0$, the distribution ν would be given by $\delta_{-\mu}$.

As for the classical Donsker theorem the Donsker theorem for deconvolution estimators has many different applications, the most obvious being the construction of confidence bands. Further Donsker theorems may be obtained by applying the functional delta method to Hadamard differentiable maps. Let us illustrate the construction of confidence bands. By the continuous mapping theorem we infer

$$\sup_{t \in \mathbb{R}} \sqrt{n} |\hat{\vartheta}_t - \vartheta_t| \xrightarrow{\mathcal{L}} \sup_{t \in \mathbb{R}} |\mathbb{G}(t)|.$$

The construction of confidence bands reduces now to knowledge about the distribution of the supremum of \mathbb{G} . Suprema of Gaussian processes are well studied and information about their distribution can be either obtained from theoretical considerations as in van der Vaart and Wellner [30, App. A.2] or from Monte Carlo simulations. Let $q_{1-\alpha}$ be the $(1-\alpha)$ -quantile of $\sup_{t \in \mathbb{R}} |\mathbb{G}(t)|$ that is $\mathbb{P}(\sup_{t \in \mathbb{R}} |\mathbb{G}(t)| \leq q_{1-\alpha}) = 1 - \alpha$. Then

$$\lim_{n \rightarrow \infty} \mathbb{P} \left(\vartheta_t \in [\hat{\vartheta}_t - q_{1-\alpha} n^{-1/2}, \hat{\vartheta}_t + q_{1-\alpha} n^{-1/2}] \text{ for all } t \in \mathbb{R} \right) = 1 - \alpha$$

and thus the intervals $[\hat{\vartheta}_t - q_{1-\alpha} n^{-1/2}, \hat{\vartheta}_t + q_{1-\alpha} n^{-1/2}]$ define a confidence band.

3 Efficiency

Having established the asymptotic normality of our estimator, the natural question is whether it is optimal in the sense of the convolution Theorem 5.2.1 by Bickel et al. [1]. Typically, efficiency is investigated for estimators T_n which are (locally) regular, that is for any parametric submodel $\eta \rightarrow f_{X,\eta}$ and $n^{1/2}|\eta_n - \eta| \lesssim 1$ the law of $n^{1/2}(T_n - \langle \zeta, f_{X,\eta} \rangle)$ under η_n converges for $n \rightarrow \infty$ to a distribution independent of (η_n) . In Lemma 9 we show that the estimator $\hat{\vartheta}_t$ from (3) is asymptotically linear with influence function $x \mapsto \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta(y)(\delta_x - \mathbb{P})(dy)$ and thus $\hat{\vartheta}_t$ is Gaussian regular.

In general, semiparametric lower bounds are constructed as the supremum of the information bounds over all regular parametric submodels. As it turns out, it suffices to apply the Cramér–Rao bound to the least favorable one-dimensional submodel \mathbb{P}_g of the form

$$f_{Y,\xi g} = f_{X,\xi g} * f_\varepsilon \quad \text{with} \quad f_{X,\xi g} := f_X + \xi g, \quad \text{for all } \xi \in (-\tau, \tau),$$

with some $\tau > 0$ and a perturbation g satisfying

$$f_X \pm \tau g \geq 0 \quad \text{and} \quad \int g = 0. \tag{8}$$

Note that all laws \mathbb{P}_g are absolutely continuous with respect to \mathbb{P} assuming $\text{supp}(f_X) = \mathbb{R}$. Moreover, the submodels are regular with score function $g * f_\varepsilon / f_Y$, since for all $\xi \in (-\tau, \tau) \setminus \{0\}$ we have the L^2 -differentiability

$$\int \left(\frac{f_{Y,\xi g} - f_Y - \xi g * f_\varepsilon}{\xi f_Y} \right)^2 f_Y = 0.$$

Similarly to van der Vaart [29, Chap. 25.5], we define the score operator $Sg := (g * f_\varepsilon) f_Y^{-1/2}$ and thus the information operator of f_X is given by $I := S^* S$, where S^* denotes the adjoint of the linear operator S . This yields the Fisher information in direction g

$$\langle I g, g \rangle = \langle S g, S g \rangle = \int \left(\frac{g * f_\varepsilon}{f_Y} \right)^2 f_Y \tag{9}$$

and we obtain the information bound

$$\mathcal{I}_\zeta := \sup_g \frac{\langle g, \zeta \rangle^2}{\langle S g, S g \rangle}, \tag{10}$$

where the supremum is taken over all g satisfying (8). In the notation of [1, Def. 3.3.2], we consider the tangent space $\dot{Q} := \{(g * f_\varepsilon)/f_Y | g \text{ satisfies (8)}\}$, representing the submodel $\{\mathbb{P}_g\}$, and the efficient influence function of the parameter $\vartheta_\zeta : \dot{Q} \rightarrow \mathbb{R}, h \mapsto \langle h, \zeta \rangle$ needs to be determined.

Since we perturb the density additively with the restriction (8), the quotient $|g/f_X|$ needs to be bounded and thus it is natural to assume a lower bound for the decay behavior of f_X . We state with some $\delta > 0$ and $M \in \mathbb{N}$

Assumption 3. *Let the following be satisfied*

- (i) f_X is bounded and fulfills the moment condition $\int |x|^{2+\delta} f_X(x) dx < \infty$,
- (ii) $f_X \in W_1^2(\mathbb{R})$ that is f_X has L^1 -Sobolev regularity two,
- (iii) $f_X(x) \gtrsim \langle x \rangle^{-M}$ for $x \in \mathbb{R}$.

A precise definition of the L^1 -Sobolev space $W_1^2(\mathbb{R})$ can be found in the appendix. Due to the Sobolev embedding $W_1^2(\mathbb{R}) \subseteq H^\alpha(\mathbb{R})$ with $\alpha < 3/2$ (cf. (44) and (46)), Assumption 3 implies the Assumption 1 in the previous section. The conditions on ε need to be strengthened, too.

Assumption 4. *We suppose*

- (i) $\int |x|^{2+\delta} f_\varepsilon(x) dx < \infty$,
- (ii) for some $\beta \in (0, \infty) \setminus \mathbb{Z}$ and M from above let $\varphi_\varepsilon \in C^{(\lfloor \beta \rfloor \vee M)+1}(\mathbb{R})$ satisfy for all $k = 0, \dots, (\lfloor \beta \rfloor \vee M) + 1$

$$\mathbb{1}_{\{k=0\}} \langle u \rangle^{-\beta-k} \lesssim |\varphi_\varepsilon^{(k)}(u)| \lesssim \langle u \rangle^{-\beta-k}.$$

Since $M + 1 \geq 2$, easy calculus shows that Assumption 2(ii) on φ_ε^{-1} follows from Assumption 4 on φ_ε . We supposed $\beta \notin \mathbb{Z}$ mainly to simplify our proofs. Let us first show an information bound for smooth ζ .

Theorem 2. *Grant Assumptions 3 and 4 and let $\zeta \in \mathcal{S}(\mathbb{R})$ be a Schwartz function. For any regular estimator T of $\vartheta_0 = \langle \zeta, f_X \rangle$ with asymptotic variance σ^2 we obtain*

$$\sigma^2 \geq \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta)^2 f_Y - \vartheta_0^2. \quad (11)$$

In particular, the supremum in (10) is attained at $g^ := g^*(\zeta) := \mathbf{I}^{-1} \zeta - \langle \zeta, f_X \rangle f_X$, where the inverses of S^* and \mathbf{I} are given by*

$$\begin{aligned} (S^*)^{-1} \zeta &= (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta) \sqrt{f_Y} \quad \text{and} \\ \mathbf{I}^{-1} \zeta &= S^{-1}(S^{-1})^* \zeta = \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * \{(\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta) f_Y\}. \end{aligned}$$

Therefore, the score function corresponding to $g^*(\zeta)$ which is given by

$$\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta - \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta) f_Y$$

(compare (37) below) is the efficient influence function and, moreover, equals the influence function of ϑ_ζ . This equality shows that the estimator is efficient for smooth functionals ϑ_ζ .

Moreover, we found already the efficient influence function in the larger tangent set of all regular submodels.

Unfortunately, less smooth ζ might be only in the domain of $(S^*)^{-1}$ while $I^{-1}\zeta$ is not in $L^2(\mathbb{R})$ and thus the formal maximizer $g^*(\zeta)$ cannot be applied rigorously as the following example shows.

Example 4. Let ε_j be gamma distributed with density $\gamma_{\beta,1}$ for $\beta \in (1/4, 1/2)$ and consider $\zeta(x) = e^x \mathbb{1}_{(-\infty, 0]}(x) = \gamma_{1,1}(-x)$ which is contained in Z^{γ_s, γ_c} for all $\gamma_s < 1/2$ and γ_c arbitrary large. We obtain

$$(S^*)^{-1}\zeta = \gamma_{1-\beta,1}(-\bullet)\sqrt{f_Y} \quad \text{and} \quad I^{-1}\zeta = \mathcal{F}^{-1}[(1-iu)^\beta((1+iu)^{-1+\beta} * \varphi)].$$

While first term behaves nicely the Fourier transform of $I^{-1}\zeta$ is of order $|u|^{-1+2\beta} > |u|^{-1/2}$ for $|u| \rightarrow \infty$ and thus $I^{-1}\zeta \notin L^2(\mathbb{R})$.

Therefore, we choose an approximating sequence $\zeta_n \rightarrow \zeta$ with $(\zeta_n)_{n \in \mathbb{N}} \subseteq \mathcal{S}(\mathbb{R})$. For $n \in \mathbb{N}$ let $g_n^* := g^*(\zeta_n) = I^{-1}\zeta_n - \langle \zeta, f_X \rangle f_X$ be the least favorable direction in the estimation problem with respect to $\langle f_X, \zeta_n \rangle$. We obtain for every $n \in \mathbb{N}$

$$\mathcal{I}_\zeta \geq \frac{\langle g_n^*, \zeta \rangle^2}{\langle Sg_n^*, Sg_n^* \rangle} = \frac{(\langle g_n^*, \zeta - \zeta_n \rangle + \langle g_n^*, \zeta_n \rangle)^2}{\langle Sg_n^*, Sg_n^* \rangle}.$$

This inequality suggests two possibilities to understand our strategy for obtaining the efficiency bound. First, the sequence (g_n^*) approximates the formal maximizer $g^*(\zeta)$ and thus plugging g_n^* into the bound (10) might converge to the supremum. Second, any unbiased estimator of $\vartheta_{\zeta_n} = \langle f_X, \zeta_n \rangle$ is at the same time a possibly biased estimator of ϑ_ζ with bias tending to zero. Therefore, the bound for the smooth problems should converge to the nonsmooth one. The following lemma provides a sufficient condition for the convergence of the Cramér–Rao bounds.

Lemma 3. *Let ζ and (ζ_n) satisfy $(S^*)^{-1}\zeta \in L^2(\mathbb{R})$ and $\zeta_n, I^{-1}\zeta_n \in L^2(\mathbb{R})$ for all $n \in \mathbb{N}$. Then $\vartheta_{\zeta_n} \rightarrow \vartheta_\zeta$ and $\frac{\langle g_n^*, \zeta \rangle^2}{\langle Sg_n^*, Sg_n^* \rangle} \rightarrow \langle (S^*)^{-1}\zeta, (S^*)^{-1}\zeta \rangle - \langle \zeta, f_X \rangle^2$ hold as $n \rightarrow \infty$ if*

$$\|(S^*)^{-1}(\zeta_n - \zeta)\|_{L^2} \rightarrow 0, \quad \text{as } n \rightarrow \infty.$$

Using mapping properties on Besov spaces, we will show that the underlying Fourier multiplier $\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}]$ and thus the inverse adjoint score operator $(S^*)^{-1}$ are well-defined on the set Z^{γ_s, γ_c} . This allows the extension of Theorem 2 to all $\zeta \in Z^{\gamma_s, \gamma_c}$ with $\gamma_s > \beta$ and $\gamma_c > \beta + 1/2$.

Since $\hat{\vartheta}_t$ does not only estimate ϑ_t pointwise but also as a process in $\ell^\infty(\mathbb{R})$, we want to generalize Theorem 2 in this direction, too. In view of Theorem 25.48 of van der Vaart [29] the remaining ingredient is the tightness of the limiting object, which is already a necessary condition for the Donsker theorem. A regular estimator T_n of $(\vartheta_t)_{t \in \mathbb{R}}$ in $\ell^\infty(\mathbb{R})$ is efficient if the limiting distribution of $\sqrt{n}(T_n - \vartheta)$ is a tight zero mean Gaussian process whose covariance structure is given by the information bound for the finite dimensional distributions (cf. the convolution Theorem 5.2.1 of [1]). Interestingly, the class of efficient influence functions for $t \in \mathbb{R}$ is not Donsker as discussed above and thus there exists no efficient estimator which is asymptotically linear in $\ell^\infty(\mathbb{R})$ [cf. 20, Thm. 18.8].

Theorem 4. *Let Assumptions 3 and 4 be satisfied as well as $\zeta \in Z^{\gamma_s, \gamma_c}$ with $\gamma_s > \beta$ and $\gamma_c > \beta + 1/2$. Then the estimator $(\hat{\vartheta}_t)_{t \in \mathbb{R}}$ defined in (3) is (uniformly) efficient.*

Additionally, the proof of Theorem 4 reveals the relation between the intrinsic metric $d(s, t)^2 = \mathbb{E}[(\mathbb{G}_s - \mathbb{G}_t)^2]$ of the limit \mathbb{G} , which is essential to show tightness, and the metric $d_{I^{-1}}(s, t)^2 = \langle (S^\star)^{-1}(\zeta_t - \zeta_s), (S^\star)^{-1}(\zeta_t - \zeta_s) \rangle$ which is induced by the inverse Fisher information, namely

$$d_{I^{-1}}(s, t)^2 = d(s, t)^2 + \langle \zeta_t - \zeta_s, f_X \rangle^2$$

(cf. equations (25) and (43) below) such that both metrics are equal up to some centering term which is another way of interpreting the efficiency of $\widehat{\vartheta}_\bullet$.

4 Proof of the Donsker theorem

First, we provide an auxiliary lemma, which describes the properties of the deconvolution operator $\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}]$.

Lemma 5. *Grant Assumption 2.*

- (i) *For all $s \in \mathbb{R}, p, q \in [1, \infty]$ the deconvolution operator $\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)]$ is a Fourier multiplier from $B_{p,q}^s(\mathbb{R})$ to $B_{p,q}^{s-\beta}(\mathbb{R})$, that is the linear map*

$$B_{p,q}^s(\mathbb{R}) \rightarrow B_{p,q}^{s-\beta}(\mathbb{R}), f \mapsto \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] \mathcal{F} f$$

is bounded.

- (ii) *For any integer m strictly larger than β we have $\mathcal{F}^{-1}[(1 + iu)^{-m} \varphi_\varepsilon^{-1}] \in L^1(\mathbb{R})$ and if $m > \beta + 1/2$ we also have $\mathcal{F}^{-1}[(1 + iu)^{-m} \varphi_\varepsilon^{-1}] \in L^2(\mathbb{R})$.*

- (iii) *Let $\beta^+ > \beta$ and $f, g \in H^{\beta^+}(\mathbb{R})$. Then*

$$\int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * f) g = \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * g) f. \quad (12)$$

Using the kernel K , this equality extends to functions $g \in L^2(\mathbb{R}) \cup L^\infty(\mathbb{R})$ and finite Borel measures μ :

$$\int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] \mathcal{F} K_h * \mu) g = \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] \mathcal{F} K_h * g) d\mu. \quad (13)$$

Proof.

- (i) Analogously to [24], we deduce from Corollary 4.11 of [16] that $(1 + iu)^{-\beta} \varphi_\varepsilon^{-1}(-u)$ is a Fourier multiplier on $B_{p,q}^s$ by Assumption 2(ii). It remains to note that $j : B_{p,q}^s(\mathbb{R}) \rightarrow B_{p,q}^{s-\beta}(\mathbb{R}), f \mapsto \mathcal{F}^{-1}[(1 + iu)^\beta \mathcal{F} f]$ is a linear isomorphism [28, Thm. 2.3.8].
- (ii) Since the gamma density $\gamma_{1,1}$ is of bounded variation, it is contained in $B_{1,\infty}^1(\mathbb{R})$ by (51). Using the isomorphism j from (i), we deduce $\gamma_{m,1} \in B_{1,\infty}^m(\mathbb{R})$ and thus by Besov embeddings (47) and (44)

$$\mathcal{F}^{-1}[(1 + iu)^{-m} \varphi_\varepsilon^{-1}] \in B_{1,\infty}^{m-\beta}(\mathbb{R}) \subseteq B_{1,1}^0(\mathbb{R}) \subseteq L^1(\mathbb{R}).$$

If $m - \beta > 1/2$ we can apply the embedding $B_{1,\infty}^{m-\beta}(\mathbb{R}) \subseteq B_{2,\infty}^{m-\beta-1/2}(\mathbb{R}) \subseteq L^2(\mathbb{R})$.

(iii) For $f \in H^{\beta+}(\mathbb{R})$ (i) and the Besov embeddings (44), (46) and (47) yield

$$\|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * f\|_{L^2} \lesssim \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * f\|_{B_{2,1}^0} \lesssim \|f\|_{B_{2,1}^\beta} \lesssim \|f\|_{H^{\beta+}} < \infty.$$

Therefore, it follows by Plancherel's equality

$$\begin{aligned} \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * f)(x)g(x) dx &= \frac{1}{2\pi} \int \varphi_\varepsilon^{-1}(-u) \mathcal{F} f(-u) \mathcal{F} g(u) du \\ &= \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * g)(x)f(x) dx. \end{aligned}$$

To prove the second part of the claim for $g \in L^2(\mathbb{R})$, we note that by Young's inequality

$$\|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1} \mathcal{F} K_h]\|_{L^2} \leq \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1} \mathbf{1}_{[-1/h, 1/h]}]\|_{L^2} \|K_h\|_{L^1} < \infty$$

due to the support of $\mathcal{F} K$ and Assumption (5) on the decay of K . Since μ is a finite measure and g is bounded, Fubini's theorem yields then

$$\begin{aligned} &\int g(x) (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1} \mathcal{F} K_h] * \mu)(x) dx \\ &= \int \int g(x) \mathcal{F}^{-1}[\varphi_\varepsilon^{-1} \mathcal{F} K_h](x-y) \mu(dy) dx \\ &= \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F} K_h] * g)(y) \mu(dy), \end{aligned}$$

where we have used the symmetry of the kernel. In order to apply Fubini's theorem for the case $g \in L^\infty(\mathbb{R})$, too, we have to show that $\|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1} \mathcal{F} K_h]\|_{L^1}$ is finite. We replace the indicator function by a function $\chi \in C^\infty(\mathbb{R})$ which equals one on $[-1/h, 1/h]$ and has got compact support. We estimate

$$\|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1} \mathcal{F} K_h]\|_{L^1} \leq \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1} \chi]\|_{L^1} \|K_h\|_{L^1}. \quad (14)$$

Using $\varphi_\varepsilon^{-1} \chi$ is twice continuously differentiable and has got compact support we obtain

$$\begin{aligned} \|(1+x^2) \mathcal{F}^{-1}[\varphi_\varepsilon^{-1} \chi](x)\|_\infty &\leq \|\mathcal{F}^{-1}[(\text{Id} - D^2) \varphi_\varepsilon^{-1} \chi](x)\|_\infty \\ &\leq \|(\text{Id} - D^2) \varphi_\varepsilon^{-1} \chi\|_{L^1} < \infty, \end{aligned}$$

where we denote the identity and the differential operator by Id and D , respectively. This shows that (14) is finite. \square

4.1 Convergence of the finite dimensional distributions

As usual, we decompose the error into a stochastic error term and a bias term:

$$\begin{aligned} \widehat{\vartheta}_t - \vartheta_t &= \widehat{\vartheta}_t - \mathbb{E}[\widehat{\vartheta}_t] + \mathbb{E}[\widehat{\vartheta}_t] - \vartheta_t \\ &= \int \zeta_t(x) \mathcal{F}^{-1} \left[\mathcal{F} K_h \frac{\varphi_n - \varphi}{\varphi_\varepsilon} \right](x) dx + \int \zeta_t(x) (K_h * f_X(x) - f_X(x)) dx. \end{aligned}$$

4.1.1 The bias

The bias term can be estimated by the standard kernel estimator argument. Let us consider the singular and the continuous part of ζ separately. Applying Plancherel's identity and Hölder's inequality, we obtain

$$\begin{aligned} & \int |\zeta_t^s(x)(K_h * f_X(x) - f_X(x))| dx \\ &= \frac{1}{2\pi} \int |\mathcal{F} \zeta_t^s(u)(\mathcal{F} K(hu) - 1) \mathcal{F} f_X(-u)| du \\ &\leq \| \langle u \rangle^{-(\alpha+\gamma_s)} (\mathcal{F} K(hu) - 1) \|_\infty \int \langle u \rangle^{\alpha+\gamma_s} |\mathcal{F} \zeta_t^s(u) \mathcal{F} f_X(u)| du \\ &\leq h^{\alpha+\gamma_s} \| u^{-(\alpha+\gamma_s)} (\mathcal{F} K(u) - 1) \|_\infty \| \zeta^s \|_{H^{\gamma_s}} \| f_X \|_{H^\alpha} \end{aligned}$$

The term $\| u^{-(\alpha+\gamma_s)} (\mathcal{F} K(u) - 1) \|_\infty$ is finite using the a Taylor expansion of $\mathcal{F} K$ around 0 with $(\mathcal{F} K)^{(l)} = 0$ for $l = 1, \dots, \lfloor \alpha + \gamma_s \rfloor$ by the order of the kernel (4).

For the smooth part of ζ_t Plancherel's identity yields

$$\begin{aligned} & \int |\zeta_t^c(x)(K_h * f_X - f_X)(x)| dx \\ &= \frac{1}{2\pi} \int |\mathcal{F} [\frac{1}{ix+1} \zeta_t^c(x)] (\text{Id} + D) \{ (\mathcal{F} K(hu) - 1) \mathcal{F} f_X(-u) \}| du \\ &\leq \int |\mathcal{F} [\frac{1}{ix+1} \zeta_t^c(x)] (\mathcal{F} K(hu) - 1 + h \mathcal{F} [ixK](hu)) \mathcal{F} f_X(-u)| du \\ &\quad - \int |\mathcal{F} [\frac{1}{ix+1} \zeta_t^c(x)] (\mathcal{F} K(hu) - 1) \mathcal{F} [ixf_X](-u)| du. \end{aligned}$$

The first term can be estimated as before and for the second term we note that $xf_X(x) \in L^2(\mathbb{R}) = H^0(\mathbb{R})$ by Assumption 1(i) such that the additional smoothness of $\frac{1}{ix+1} \zeta^c(x)$ yields the right order. Therefore, we have $|\mathbb{E}[\hat{\vartheta}_t] - \vartheta_t| \lesssim h^{\alpha+\gamma_s}$ and thus by the choice of h , the bias term is of order $o(n^{-1/2})$.

4.1.2 The stochastic error

We notice that $\| \zeta^c - a \|_{H^{\gamma_c}} \lesssim \| \langle x \rangle^{-\tau} \|_{C^s} \| \langle x \rangle^\tau (\zeta^c(x) - a(x)) \|_{H^{\gamma_c}} < \infty$ for any $s > \gamma_c$, where we used the pointwise multiplier property (48) as well as the Besov embeddings (47) and (45). We have $\zeta^s \in L^2$ and by (44), (46) and (47)

$$\| \zeta^c \|_\infty \leq \| a \|_\infty + \| \zeta^c - a \|_\infty \leq \| a \|_\infty + \| \zeta^c - a \|_{H^{\gamma_c}} < \infty,$$

since $\gamma_c > 1/2$. Consequently we can apply the smoothed adjoint equality (13) and obtain for the stochastic error term

$$\begin{aligned} & \int \zeta_t(x) \mathcal{F}^{-1} \left[\mathcal{F} K_h \frac{\varphi_n - \varphi}{\varphi_\varepsilon} \right] (x) dx \\ &= \int \mathcal{F}^{-1} [\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F} K_h] * \zeta_t(x) (\mathbb{P}_n - \mathbb{P})(dx). \end{aligned} \tag{15}$$

Therefore, it suffices for the convergence of the finite dimensional distributions to bound the term

$$\sup_{h \in (0,1)} \int |\mathcal{F}^{-1} [\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F} K_h] * \zeta(x)|^{2+\delta} \mathbb{P}(dx), \tag{16}$$

for any function $\zeta \in Z^{\gamma_s, \gamma_c}$. Then the stochastic error term converges in distribution to a normal random variable by the central limit theorem under the Lyapunov condition [i.e., 19, Thm. 15.43 together with Lem. 15.41]. Finally, the Cramér-Wold device yields the convergence of the finite dimensional distributions in Theorem 1.

First, note that the moment conditions in Assumptions 1 and 2 and the estimate

$$|x|^p f_Y(x) \leq \int |x - y + y|^p f_X(x - y) f_\varepsilon(y) dy \lesssim (|y|^p f_X) * f_\varepsilon + f_X * (|y|^p f_\varepsilon),$$

for $x \in \mathbb{R}$, $p \geq 1$, yield finite $(2 + \delta)$ th moments for \mathbb{P} since

$$\int |x|^{2+\delta} f_Y(x) dx \lesssim \| |x|^{2+\delta} f_X \|_{L^1} \| f_\varepsilon \|_{L^1} + \| f_X \|_{L^1} \| |x|^{2+\delta} f_\varepsilon \|_{L^1} < \infty. \quad (17)$$

To estimate (16), we rewrite

$$\begin{aligned} \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta^c(x) &= \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u)(\text{Id} + D) \mathcal{F}[\frac{1}{iy+1} \zeta^c(y)](u)](x) \\ &= \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F}[\frac{1}{iy+1} \zeta^c(y)](u)](x) \\ &\quad + \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) (\mathcal{F}[\frac{1}{iy+1} \zeta^c(y)])'(u)](x) \\ &= (1 + ix) \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F}[\frac{1}{iy+1} \zeta^c(y)](u)](x) \\ &\quad + \mathcal{F}^{-1}[(\varphi_\varepsilon^{-1})'(-u) \mathcal{F}[\frac{1}{iy+1} \zeta^c(y)](u)](x), \end{aligned} \quad (18)$$

owing to the product rule for differentiation. Hence,

$$\begin{aligned} \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta(x) &= \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta^s(u)](x) \\ &\quad + (1 + ix) \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F}[\frac{1}{iy+1} \zeta^c(y)](u)](x) \\ &\quad + \mathcal{F}^{-1}[(\varphi_\varepsilon^{-1})'(-u) \mathcal{F}[\frac{1}{iy+1} \zeta^c(y)](u)](x). \end{aligned} \quad (19)$$

While $\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta$ may exist only in distributional sense in general, it is defined rigorously through the right-hand side of the above display for $\zeta \in Z^{\gamma_s, \gamma_c}$. Considering $\zeta * K_h$ instead of ζ , we estimate separately all three terms in the following.

The continuity and linearity of the Fourier multiplier $\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)]$, which was shown in Lemma 5(i), yield for the first term in (19)

$$\begin{aligned} \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta^s(u) \mathcal{F} K_h(u)]\|_{H^\delta} &= \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F}[\zeta^s * K_h]]\|_{B_{2,2}^\delta} \\ &\lesssim \|\zeta^s * K_h\|_{B_{2,2}^{\beta+\delta}} \lesssim \|\zeta^s\|_{H^{\beta+\delta}}, \end{aligned}$$

where the last inequality holds by $\|\mathcal{F} K_h\|_\infty \leq \|K\|_{L^1}$. Using the boundedness of f_Y and the continuous Sobolev embedding $H^{\delta/4}(\mathbb{R}) \subseteq L^{2+\delta}(\mathbb{R})$ by (44), (47) and (46), we obtain

$$\begin{aligned} \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta^s(u) \mathcal{F} K_h(u)]\|_{L^{2+\delta}(\mathbb{P})} &\lesssim \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta^s(u) \mathcal{F} K_h(u)]\|_{L^{2+\delta}} \\ &\lesssim \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta^s(u) \mathcal{F} K_h(u)]\|_{H^\delta} \\ &\lesssim \|\zeta^s\|_{H^{\beta+\delta}} \end{aligned} \quad (20)$$

To estimate the second term in (19), we use the Cauchy–Schwarz inequality and Assumption 2(ii):

$$\begin{aligned}
& \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F}[\frac{1}{ix+1}\zeta^c(x)](u) \mathcal{F}K_h(u)]\|_\infty \\
& \leq \|\varphi_\varepsilon^{-1}(-u) \mathcal{F}[\frac{1}{ix+1}\zeta^c] \mathcal{F}K_h(u)\|_{L^1} \\
& \lesssim \|\langle u \rangle^{-1/2-\beta-\delta} \varphi_\varepsilon^{-1}(-u)\|_{L^2} \|\langle u \rangle^{1/2+\beta+\delta} \mathcal{F}[\frac{1}{ix+1}\zeta^c(x)]\|_{L^2} \\
& \lesssim \|\frac{1}{ix+1}\zeta^c(x)\|_{H^{1/2+\beta+\delta}}.
\end{aligned}$$

Thus $\int (1+x^2)^{(2+\delta)/2} f_Y(x) dx < \infty$ from (17) yields

$$\begin{aligned}
& \|(1+ix) \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F}[\frac{1}{iy+1}\zeta^c(y)](u) \mathcal{F}K_h(u)](x)\|_{L^{2+\delta}(\mathbb{P})} \\
& \lesssim \|\frac{1}{ix+1}\zeta^c(x)\|_{H^{1/2+\beta+\delta}}.
\end{aligned} \tag{21}$$

The last term in the decomposition (19) can be estimated similarly using the Cauchy–Schwarz inequality and Assumption 2(ii) for $(\varphi^{-1})'$

$$\begin{aligned}
& \|\mathcal{F}^{-1}[(\varphi_\varepsilon^{-1})'(-u) \mathcal{F}[\frac{1}{ix+1}\zeta^c(x)](u) \mathcal{F}K_h(u)]\|_{L^{2+\delta}(\mathbb{P})} \\
& \lesssim \|(\varphi_\varepsilon^{-1})'(-u) \mathcal{F}[\frac{1}{ix+1}\zeta^c(x)](u)\|_{L^1} \\
& \leq \|\langle u \rangle^{1/2-\beta-\delta} (\varphi_\varepsilon^{-1})'\|_{L^2} \|\langle u \rangle^{-1/2+\beta+\delta} \mathcal{F}^{-1}[\frac{1}{ix+1}\zeta^c(x)](u)\|_{L^2} \\
& \lesssim \|\frac{1}{ix+1}\zeta^c(x)\|_{H^{-1/2+\beta+\delta}}.
\end{aligned} \tag{22}$$

Combining (20), (21) and (22), we obtain

$$\sup_{h \in (0,1)} \|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F}K_h] * \zeta(x)\|_{L^{2+\delta}(\mathbb{P})} \lesssim \|\zeta\|_{Z^{\beta+\delta, 1/2+\beta+\delta}}, \tag{23}$$

which is finite for δ small enough satisfying $\beta + \delta \leq \gamma_s$ and $1/2 + \beta + \delta \leq \gamma_c$. Since $\mathcal{F}K_h$ converges pointwise to one and $|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F}K_h] * \zeta(x)|^2$ is uniformly integrable by the bound of the $2 + \delta$ moments, the variance converges to

$$\int |\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta(x)|^2 \mathbb{P}(dx).$$

4.2 Tightness

Motivated by the representation (15) of the stochastic error, we introduce the empirical process

$$\nu_n(t) := \sqrt{n} \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F}K_h] * \zeta_t(x) (\mathbb{P}_n - \mathbb{P})(dx), \quad t \in \mathbb{R}. \tag{24}$$

In order to show tightness of the empirical process, we first show some properties of the class of translations $\mathcal{H} := \{\zeta_t | t \in \mathbb{R}\}$ for $\zeta \in Z^{\gamma_s, \gamma_c}$.

Lemma 6. *For $\zeta \in Z^{\gamma_s, \gamma_c}$ the following is satisfied:*

- (i) *The decomposition $\zeta_t = \zeta_t^c + \zeta_t^s$ satisfies the conditions in the definition of Z^{γ_s, γ_c} with a_t . We have $\sup_{t \in \mathbb{R}} \|\zeta_t\|_{Z^{\gamma_s, \gamma_c}} < \infty$.*

(ii) For any $\eta \in (0, \gamma_s)$ there is a $\tau > 0$ such that $\|\zeta_t - \zeta_s\|_{Z^{\gamma_s-\eta, \gamma_c-\eta}} \lesssim |t-s|^\tau$ holds for all $s, t \in \mathbb{R}$ with $|t-s| \leq 1$

Proof.

(i) Since $\|\zeta_t^s\|_{H^{\gamma_s}}^2 = \int \langle u \rangle^{2\gamma_s} |e^{itu} \mathcal{F} \zeta^s(u)|^2 du = \|\zeta^s\|_{H^{\gamma_s}}^2$, both claims hold for the singular part. Applying the pointwise multiplier property of Besov spaces (48) as well as the Besov embeddings (47) and (45), we obtain for some $M > \gamma_c$ and $a \in C^\infty(\mathbb{R})$ as in definition (6)

$$\begin{aligned} \|\langle x \rangle^\tau (\zeta_t^c(x) - a_t(x))\|_{H^{\gamma_c}} &\lesssim \|\frac{\langle x \rangle^\tau}{\langle x-t \rangle^\tau}\|_{C^M} \|\langle x-t \rangle^\tau (\zeta_t^c(x) - a_t(x))\|_{H^{\gamma_c}} \\ &= \|\frac{\langle x \rangle^\tau}{\langle x-t \rangle^\tau}\|_{C^M} \|\langle x \rangle^\tau (\zeta^c(x) - a(x))\|_{H^{\gamma_c}}, \end{aligned}$$

which is finite for all $t \in \mathbb{R}$ since $\langle x \rangle^\tau \langle x-t \rangle^{-\tau} \in C^M(\mathbb{R})$. For the second claim we estimate similarly

$$\begin{aligned} \sup_{t \in \mathbb{R}} \|\frac{1}{ix+1} \zeta_t^c(x)\|_{H^{\gamma_c}} &\lesssim \sup_{t \in \mathbb{R}} \|\frac{a_t(x)}{ix+1}\|_{H^{\gamma_c}} + \|\frac{1}{ix+1}\|_{C^M} \sup_{t \in \mathbb{R}} \|\zeta_t^c - a_t\|_{H^{\gamma_c}} \\ &\lesssim \|\frac{1}{ix+1}\|_{H^{\gamma_c}} \|a\|_{C^M} + \|\frac{1}{ix+1}\|_{C^M} \|\zeta^c - a\|_{H^{\gamma_c}} < \infty. \end{aligned}$$

(ii) For the singular part note that

$$\begin{aligned} &\|\zeta_t^s - \zeta_s^s\|_{H^{\gamma_s-\eta}} \\ &\leq \|\langle u \rangle^{\gamma_s} \mathcal{F} \zeta^s(u)\|_{L^2} \|\langle u \rangle^{-\eta} (1 - e^{i(t-s)u})\|_\infty \\ &\lesssim \|\langle u \rangle^{-\eta}\|_{L^\infty(\mathbb{R} \setminus (-|t-s|^{-1/2}, |t-s|^{-1/2}))} \\ &\quad \vee \|(1 - e^{i(t-s)u})\|_{L^\infty((-|t-s|^{-1/2}, |t-s|^{-1/2}))} \\ &\lesssim |t-s|^{\eta/2} \vee |t-s|^{1/2}. \end{aligned}$$

For ζ^c we have

$$\begin{aligned} \left\| \frac{1}{ix+1} (\zeta_t^c(x) - \zeta_s^c(x)) \right\|_{H^{\gamma_c-\eta}} &\lesssim \left\| \frac{1}{ix+1} \zeta_t^c(x) - \left(\frac{1}{ix+1} \zeta_t^c(x) \right) * \delta_{s-t} \right\|_{H^{\gamma_c-\eta}} \\ &\quad + \left\| \frac{1}{i(x-s+t)+1} \zeta_s^c(x) - \frac{1}{ix+1} \zeta_s^c(x) \right\|_{H^{\gamma_c-\eta}}. \end{aligned}$$

The first term can be treated analogously to ζ^s . Using some integer $M \in \mathbb{N}$ strictly larger than γ_c , the second term can be estimated by

$$\begin{aligned} &\left\| \frac{1}{i(x-s+t)+1} \zeta_s^c(x) - \frac{1}{ix+1} \zeta_s^c(x) \right\|_{H^{\gamma_c-\eta}} \\ &\lesssim |t-s| \left\| \frac{1}{i(x-s+t)+1} \frac{1}{ix+1} \zeta_s^c(x) \right\|_{H^{\gamma_c-\eta}} \\ &\lesssim |t-s| \left\| \frac{1}{i(x-s+t)+1} \right\|_{C^M} \left\| \frac{1}{ix+1} \zeta_s^c(x) \right\|_{H^{\gamma_c-\eta}} \\ &\lesssim |t-s|, \end{aligned}$$

where we used again pointwise multiplier (48), embedding properties of Besov spaces (47) and (45) as well as (i). \square

4.2.1 Pregaussian limit process

Let \mathbb{G} be the stochastic process from Theorem 1. It induces the intrinsic covariance metric $d(s, t) := \mathbb{E}[(\mathbb{G}_s - \mathbb{G}_t)^2]^{1/2}$.

Theorem 7. *There exists a version of \mathbb{G} with uniformly d -continuous sample paths almost surely and with $\sup_{t \in \mathbb{R}} |\mathbb{G}_t| < \infty$ almost surely.*

The proof of the theorem shows in addition that \mathbb{R} is totally bounded with respect to d . The boundedness of the sample paths follows from the totally bounded index set and the uniform continuity. Further we conclude that \mathcal{G} defined in (7) is \mathbb{P} -pregaussian by van der Vaart and Wellner [30, p. 89]. Thus \mathbb{G} is a tight Borel random variable in $\ell^\infty(\mathbb{R})$ and the law of \mathbb{G} is uniquely defined through the covariance structure and the sample path properties in the theorem [30, Lem. 1.5.3].

Proof. To show that the class is pregaussian, it suffices to verify polynomial covering numbers. To that end we deduce that

$$d(s, t) = (\|g_t - g_s\|_{L^2(\mathbb{P})}^2 - \langle \zeta_t - \zeta_s, f_X \rangle^2)^{1/2} \leq \|g_t - g_s\|_{L^2(\mathbb{P})} \quad (25)$$

decreases polynomial for $|t - s| \rightarrow 0$, for $\max(s, t) \rightarrow \infty$ and for $\min(s, t) \rightarrow \infty$. Using the same estimates which show the moment bound (23) but replacing $\mathcal{F}K_h = 1$, we obtain

$$\|\mathcal{F}[\varphi_\epsilon^{-1}(-\bullet)] * \zeta\|_{L^2(\mathbb{P})} \lesssim \|\zeta\|_{Z^{\beta+\delta, 1/2+\beta+\delta}} \quad (26)$$

and thus by choosing δ and η small enough Lemma 6 yields $d(s, t) \lesssim \|\zeta_t - \zeta_s\|_{Z^{\beta+\delta, 1/2+\beta+\delta}} \lesssim |t - s|^\tau$. We now turn to the estimation of the tails. We will only consider the case $s, t \geq N$ since the case $s, t \leq N$ can be treated in the same way. Without loss of generality, let $s < t$.

For the smooth component of ζ we have to show that $\|\frac{1}{ix+1}(\zeta_t^c(x) - \zeta_s^c(x))\|_{H^{\gamma_c}}$ with $t, s \geq N$ decays polynomially in N . It suffices to prove $\|\frac{1}{ix+1}(\zeta_t^c - a_t)(x)\|_{H^{\gamma_c}}$ and $\|\frac{1}{ix+1}(a_t - a_s)(x)\|_{H^{\gamma_c}}$ with $a \in C^\infty(\mathbb{R})$ from definition (6) of Z^{γ_s, γ_c} both decay polynomially in N . Let $M > \gamma_c$ and $\psi \in C^M(\mathbb{R})$ with $\psi(x) = 1$ for $x \in \mathbb{R} \setminus [-\frac{1}{2}, \frac{1}{2}]$ and $\psi(x) = 0$ for $x \in [-\frac{1}{4}, \frac{1}{4}]$. The pointwise multiplier property (48) yields

$$\begin{aligned} & \left\| \frac{1}{ix+1}(\zeta_t^c - a_t)(x) \right\|_{H^{\gamma_c}} \\ &= \left\| (\psi(x/N) + (1 - \psi(x/N))) \frac{1}{ix+it+1}(\zeta^c - a)(x) \right\|_{H^{\gamma_c}} \\ &\lesssim \left\| \frac{1}{ix+it+1} \right\|_{C^M} \|\psi(x/N)(\zeta^c - a)(x)\|_{H^{\gamma_c}} + \left\| \frac{1 - \psi(x/N)}{ix+it+1} \right\|_{C^M} \|\zeta^c - a\|_{H^{\gamma_c}} \\ &\lesssim \|\langle x \rangle^{-\tau} \psi(x/N)\|_{C^M} \|\langle x \rangle^\tau (\zeta^c - a)(x)\|_{H^{\gamma_c}} + N^{-1} \|\zeta^c - a\|_{H^{\gamma_c}} \\ &\lesssim N^{-(\tau \wedge 1)} \end{aligned}$$

and for N large enough such that $\text{supp}(a') \subseteq [-N/2, N/2]$ we obtain

$$\begin{aligned} & \left\| \frac{1}{ix+1}(a_t - a_s)(x) \right\|_{H^{\gamma_c}} \\ &= \left\| \frac{\psi(x/N)}{ix+1}(a_t - a_s)(x) \right\|_{H^{\gamma_c}} \lesssim \left\| \frac{\psi(x/N)}{ix+1} \right\|_{H^{\gamma_c}} \|(a_t - a_s)(x)\|_{C^M} \\ &\lesssim \|(ix+1)^{-3/4}\|_{H^{\gamma_c}} \|\psi(x/N)(ix+1)^{-1/4}\|_{C^M} \lesssim N^{-1/4}. \end{aligned}$$

To bound the singular part it suffices to show that

$$\|\mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet)] * \zeta_t^s\|_{L^2(\mathbb{P})}, \quad t \geq N,$$

decays polynomially in N . To this end, we split the integral domain into

$$\begin{aligned} \|\mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet)] * \zeta_t^s\|_{L^2(\mathbb{P})}^2 &= \int_{-\infty}^{-N/2} |\mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x)|^2 f_Y(x+t) dx \\ &\quad + \int_{-N/2}^{\infty} |\mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x)|^2 f_Y(x+t) dx. \end{aligned} \quad (27)$$

To estimate the first term, we use the following auxiliary calculations

$$\begin{aligned} ix \mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x) \\ = -\mathcal{F}^{-1}[(\varphi_\epsilon^{-1})'(-\bullet) \mathcal{F} \zeta^s](x) + \mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F}[iy\zeta^s(y)]](x) \end{aligned}$$

and with an integer $M \in \mathbb{N}$ strictly larger than γ_s and a function $\chi \in C^M(\mathbb{R})$ which is equal to one on $\text{supp}(\zeta^s)$ and has compact support

$$\begin{aligned} \|y\zeta^s(y)\|_{H^{\gamma_s}} &= \|y\chi(y)\zeta^s(y)\|_{H^{\gamma_s}} \lesssim \|y\chi(y)\|_{B_{\infty,2}^{\gamma_s}} \|\zeta^s(y)\|_{H^{\gamma_s}} \\ &\lesssim \|y\chi(y)\|_{C^M} < \infty, \end{aligned}$$

where we used the pointwise multiplier property (48) of Besov spaces as well as the Besov embeddings (47) and (45). Thus $ix \mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x) \in L^2(\mathbb{R})$. Applying this and the boundedness of f_Y to the first term in (27) yields

$$\begin{aligned} &\int_{-\infty}^{-N/2} |\mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x)|^2 f_Y(x+t) dx \\ &\lesssim \int_{-\infty}^{-N/2} |\mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x)|^2 dx \\ &\leq 4N^{-2} \int_{-\infty}^{-N/2} |x \mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x)|^2 dx \lesssim N^{-2}. \end{aligned}$$

Using Hölders's inequality and the boundedness of f_Y , we estimate the second term in (27) by

$$\begin{aligned} &\|\mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x)\|_{L^{2+\delta}}^2 \left(\int_{-N/2}^{\infty} |f_Y(x+t)|^{(2+\delta)/\delta} dx \right)^{\delta/(2+\delta)} \\ &\lesssim \|\mathcal{F}^{-1}[\varphi_\epsilon^{-1}(-\bullet) \mathcal{F} \zeta^s](x)\|_{L^{2+\delta}}^2 \left(\int_{N/2}^{\infty} f_Y(x) dx \right)^{\delta/(2+\delta)}. \end{aligned}$$

While the first factor is finite according to our bound (20), which also holds when $\mathcal{F} K_h$ is omitted, the second one is of order $N^{-\delta}$ due to the finite $(2+\delta)$ th moment of \mathbb{P} . Therefore, the second term in (27) decays polynomially. \square

4.2.2 Uniform central limit theorem

We recall the definition of the empirical process ν_n in (24).

Theorem 8. *Grant Assumptions 1 and 2. Let*

$$(\nu_n(t_1), \dots, \nu_n(t_k)) \xrightarrow{\mathcal{L}} (\mathbb{G}_{t_1}, \dots, \mathbb{G}_{t_k})$$

for all $t_1, \dots, t_k \in \mathbb{R}$ and for all $k \in \mathbb{N}$. If either $\gamma_s \leq \beta + 1/2$ and $h_n^\rho n^{1/4} \rightarrow \infty$ as $n \rightarrow \infty$ for some $\rho > \beta - \gamma_s + 1/2$ or if $\gamma_s > \beta + 1/2$, then

$$\nu_n \xrightarrow{\mathcal{L}} \mathbb{G} \quad \text{in } \ell^\infty(\mathbb{R}).$$

Proof. We split the empirical process ν_n into three parts

$$\nu_n = \sqrt{n} \int (T_1(x) + T_2(x) + T_3(x))(\mathbb{P}_n - \mathbb{P})(dx),$$

where T_1 , T_2 and T_3 correspond to the three terms in decomposition (19) and are given by (28), (29) and (30) below. For the first term

$$T_1(x) = \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta_t^s(u) \mathcal{F} K_h(u)](x) \quad (28)$$

we distinguish the two cases $\gamma_s > \beta + 1/2$ and $\gamma_s \leq \beta + 1/2$. In the first case we will show that T_1 varies in a fixed Donsker class. In the second case the process indexed by T_1 is critical, this is where smoothed empirical processes and the condition on the bandwidth are needed. Tightness of T_1 in this case will be shown in Section 4.2.3. We will further show that the second term T_2 and the third term T_3 are both varying in fixed Donsker classes for all $\gamma_s > \beta$. In particular the three processes indexed by T_1 , T_2 and T_3 , respectively, are tight. Applying the equicontinuity characterization of tightness [30, Thm. 1.5.7] with the maximum of the semimetrics yields that ν_n is tight. Since we have assumed convergence of the finite dimensional distribution the convergence of ν_n in distribution follows [30, Thm. 1.5.4].

Here we consider only the first case $\gamma_s > \beta + 1/2$. We recall that ζ_t^s is contained in $H^{\gamma_s}(\mathbb{R})$. By the Fourier multiplier property of the deconvolution operator in Lemma 5(i) and by $\sup_{h>0, u} |\mathcal{F} K_h(u)| \leq \|K\|_{L^1} < \infty$ the functions T_1 are contained in a bounded set of $H^{1/2+\eta}(\mathbb{R})$ for some $\eta > 0$ small enough. We apply [23, Prop. 1] with $p = q = 2$ and $s = 1/2 + \eta$ and conclude that T_1 varies in a universal Donsker class.

The second term is of the form

$$T_2(x) = (1 + ix) \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F}[\frac{1}{iy+1} \zeta_t^c(y)](u) \mathcal{F} K_h(u)](x). \quad (29)$$

By Assumption 2(ii) we have $\varphi_\varepsilon^{-1}(u) \lesssim \langle u \rangle^\beta$. For some $\eta > 0$ sufficiently small, the functions $\frac{1}{iy+1} \zeta_t^c(y)$, $t \in \mathbb{R}$, are contained in a bounded set of $H^{\beta+\eta+1/2}(\mathbb{R})$ by Lemma 6. We obtain that the functions $T_2(x)/(1+ix)$ are contained in a bounded subset of $H^{1/2+\eta}(\mathbb{R})$. Corollary 5 in [23] yields with $p = q = 2$, $\beta = -1$, $s = 1/2 + \eta$ and $\gamma = \eta$ that T_2 is contained in a fixed \mathbb{P} -Donsker class.

Similarly, we treat the third term

$$T_3(x) = \mathcal{F}^{-1}[(\varphi_\varepsilon^{-1})'(-u) \mathcal{F}[\frac{1}{iy+1} \zeta_t^c(y)](u) \mathcal{F} K_h(u)](x). \quad (30)$$

By Assumption 2(ii) we have $(\varphi_\varepsilon^{-1})' \lesssim \langle u \rangle^{\beta-1}$. As above we conclude that the functions T_3 are contained in a bounded set of $H^{\eta+3/2}(\mathbb{R})$. By [23, Prop. 1] with $p = q = 2$ and $s = \eta + 3/2$ the term T_3 varies in a universal Donsker class. \square

4.2.3 The critical term

In this section, we treat the first term T_1 in the case $\gamma_s \leq \beta + 1/2$. We define

$$q_t := \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta_t^s(u)]. \quad (31)$$

For simplicity in point (e) below it will be convenient to work with functions K_h of bounded support. Thus we fix $\xi > 0$ and define the truncated kernel

$$K_h^{(0)} := K_h \mathbf{1}_{[-\xi, \xi]}.$$

By the Assumption (5) on the decay of K we have $\sup_{h>0} \|K_h - K_h^{(0)}\|_{BV} < \infty$. We conclude $\mathcal{F}(K_h - K_h^{(0)})(u) \lesssim (1 + |u|)^{-1}$ with a constant independent of $h > 0$. By Assumption 2(ii) we have $|\varphi_\varepsilon^{-1}(u)| \lesssim (1 + |u|)^\beta$. The functions $\zeta_t^s(u)$, $t \in \mathbb{R}$, are contained in a bounded set of $H^{\gamma_s}(\mathbb{R})$. Consequently T_1 with $K_h - K_h^{(0)}$ instead of K_h is contained in a bounded set of $H^{\gamma_s - \beta + 1}(\mathbb{R})$. With the same argument as used for T_3 we see that this term is contained in a universal Donsker class because $\gamma_s - \beta + 1 > 1$ by assumption. So it remains to consider T_1 with the truncated kernel $K_h^{(0)}$.

In order to show tightness of the process indexed by T_1 with the truncated kernel $K_h^{(0)}$ we check the assumptions of Theorem 3 by Giné and Nickl [14] in the version of Nickl and Reiß [24, Thm. 12] for the class $Q = \{q_t | t \in \mathbb{R}\}$ and for $\mu_n(dx) := K_{h_n}^{(0)}(x) dx$, where $q_t(x)$ was defined in (31). By Section 4.2.1 the class \mathcal{G} is \mathbb{P} -pregaussian. From the proof also follows that Q is \mathbb{P} -pregaussian since this is just the case $\zeta^c = 0$.

We write

$$Q'_\tau := \{r - q | r, q \in Q, \|r - q\|_{L^2(\mathbb{P})} \leq \tau\}.$$

Let $\rho > \beta - \gamma_s + 1/2 \geq 0$ be such that $h_n^\rho n^{1/4} \rightarrow \infty$. We fix some $\rho' \in (\beta - \gamma_s + 1/2, \rho \wedge 1)$ and obtain $h_n^{\rho'} \log(n)^{-1/2} n^{1/4} \rightarrow \infty$. We need to verify the following conditions.

- (a) We will show that the functions in $\tilde{Q}_n := \{q_t * \mu_n | t \in \mathbb{R}\}$ are bounded by $M_n := Ch_n^{-\rho'}$ for some constant $C > 0$. Since q_t is only a translation of q_0 it suffices to consider q_0 . By the definition of Z^{γ_s, γ_c} in (6), by Lemma 5(i) and by the Besov embedding (47)

$$q_0 = \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-u) \mathcal{F} \zeta^s(u)] \in B_{2,2}^{\gamma_s - \beta}(\mathbb{R}) \subseteq B_{2,\infty}^{1/2 - \rho'}(\mathbb{R}).$$

By our assumptions on the kernel (5) it follows that K' is integrable and thus that K is of bounded variation. Next we apply continuous embeddings for Besov spaces (44) and (46), (49) as well as the estimate for $\|K_{h_n}\|_{B_{1,1}^{\rho'}}$ in Giné and Nickl [14, p. 384], which also applies to truncated kernels, and obtain

$$\|q_0 * K_{h_n}^{(0)}\|_\infty \lesssim \|q_0 * K_{h_n}^{(0)}\|_{B_{\infty,1}^0} \lesssim \|q_0 * K_{h_n}^{(0)}\|_{B_{2,1}^{1/2}} \lesssim \|K_{h_n}^{(0)}\|_{B_{1,1}^{\rho'}} \lesssim h_n^{-\rho'}. \quad (32)$$

- (b) For $r \in Q'_\tau$ holds $\|r * K_h^{(0)}\|_{L^2(\mathbb{P})} \leq \|r * K_h^{(0)} - r\|_{L^2(\mathbb{P})} + \tau$. Thus it suffices to show that $\|q * K_h^{(0)} - q\|_{L^2(\mathbb{P})} \rightarrow 0$ uniformly over $q \in Q$. We estimate

$$\|q_t * K_h^{(0)} - q_t\|_{L^2(\mathbb{P})} \lesssim \|\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F} \zeta^s(\mathcal{F} K_h^{(0)} - 1)\|_{L^2}.$$

$\varphi_\varepsilon^{-1}(-\bullet) \mathcal{F} \zeta^s$ is an L^2 -function and $\mathcal{F} K_h^{(0)}$ is uniformly bounded and converges to one as $h \rightarrow 0$. By dominated convergence the integral converges to zero.

- (c) The estimates in (a) can be used to see that the classes \tilde{Q}_n have polynomial $L^2(\mathbb{Q})$ -covering numbers, uniformly in all probability measures \mathbb{Q} and uniformly in n . The function $q_0 * K_{h_n}^{(0)}$ is the convolution of two L^2 -functions and thus continuous. The estimate (32) and embedding (50) yield that $q_0 * K_{h_n}^{(0)}$ is of finite 2-variation. We argue as in Lemma 1 by Giné and Nickl [15]. As function of bounded 2-variation $q_0 * K_{h_n}^{(0)}$ can be written as a composition $g_n \circ f_n$ of a nondecreasing function f_n and a function g_n , which satisfies a Hölder condition $|g_n(u) - g_n(v)| \leq |u - v|^{1/2}$, see, for example, [8, p. 1971]. More precisely, we can take $f_n(x)$ to be the 2-variation of $q_0 * K_{h_n}^{(0)}$ up to x and the envelopes of f_n to be multiples of $M_n^2 = C^2 h_n^{-2\rho'}$. The set F_n of all translates of the nondecreasing function f_n has VC-index 2 and thus polynomial $L^1(\mathbb{Q})$ -covering numbers [7, Thm. 5.1.15]. Since each ϵ^2 -covering of translates of f_n for $L^1(\mathbb{Q})$ induces an ϵ -covering of translates of $g_n \circ f_n$ for $L^2(\mathbb{Q})$ we can estimate the covering numbers by

$$N(\tilde{Q}_n, L^2(\mathbb{Q}), \epsilon) \leq N(F_n, L^1(\mathbb{Q}), \epsilon^2) \lesssim (M_n/\epsilon)^4,$$

with constants independent of n and \mathbb{Q} . The conditions for inequality (22) by Giné and Nickl [14] are fulfilled, where the envelopes are $M_n = C h_n^{-\rho'}$ and $H_n(\eta) = H(\eta) = C_1 \log(\eta) + C_0$ with $C_0, C_1 > 0$. Consequently

$$\mathbb{E}^* \left\| \frac{1}{\sqrt{n}} \sum_{j=1}^n \varepsilon_j f(X_j) \right\|_{(\tilde{Q}_n)'_{n^{-1/4}}} \lesssim \max \left(\frac{\sqrt{\log(n)}}{n^{1/4}}, \frac{h_n^{-\rho'}}{\sqrt{n}} \log(n) \right) \rightarrow 0$$

as $n \rightarrow \infty$.

- (d) We apply Lemma 1 of [14] to show that

$$\cup_{n \geq 1} \tilde{Q}_n = \bigcup_{n \geq 1} \left\{ x \mapsto \int_{\mathbb{R}} q_t(x - y) K_{h_n}^{(0)}(y) dy \mid t \in \mathbb{R} \right\}$$

is in the $L^2(\mathbb{P})$ -closure of $\|K\|_{L^1}$ -times the symmetric convex hull of the pregaussian class Q . The condition $q_t(\bullet - y) \in L^2(\mathbb{P})$ is satisfied for all $y \in \mathbb{R}$ since $q_t \in L^2(\mathbb{R})$ and f_Y is bounded. $q_t(x - \bullet) \in L^1(|\mu_n|)$ is fulfilled owing to $K_{h_n}^{(0)}, q_t \in L^2(\mathbb{R})$. The third condition that $y \mapsto \|q_t(\bullet - y)\|_{L^2(\mathbb{P})}$ is in $L^1(|\mu_n|)$ holds likewise since f_Y is bounded and $K_{h_n}^{(0)} \in L^1(\mathbb{R})$.

- (e) The $L^2(\mathbb{P})$ -distance of two functions in \tilde{Q}_n can be estimated by

$$\begin{aligned} & \mathbb{E} \left[(q_t * K_h^{(0)}(X) - q_s * K_h^{(0)}(X))^2 \right]^{1/2} \\ &= \left\| \int q_t(\bullet - u) K_h^{(0)}(u) - q_s(\bullet - u) K_h^{(0)}(u) du \right\|_{L^2(\mathbb{P})} \\ &\leq \int |K_h^{(0)}(u)| \|q_t(\bullet - u) - q_s(\bullet - u)\|_{L^2(\mathbb{P})} du \\ &\leq \|K_h^{(0)}\|_{L^1} \sup_{|u| \leq \xi} \|q_t(\bullet - u) - q_s(\bullet - u)\|_{L^2(\mathbb{P})} \\ &= \|K_h^{(0)}\|_{L^1} \sup_{|u| \leq \xi} \|q_{t+u} - q_{s+u}\|_{L^2(\mathbb{P})}. \end{aligned}$$

As seen in the proof that Q is pregaussian, the covering numbers grow at most polynomially. We take N large enough such that $N \geq 2\xi$. Then $s, t > N$ implies $s + u, t + u > N/2$ and $s, t < -N$ implies $s + u, t + u < -N/2$. Since this is only a polynomial change in N , the growth of the covering numbers remains at most polynomial. This leads to the entropy bound $H(\tilde{Q}_n, L^2(\mathbb{P}), \eta) \lesssim \log(\eta^{-1})$ for η small enough and independent of n . We define $\lambda_n(\eta) := \log(\eta^{-1})\eta^2$. The bound in the condition is of the order $\log(n)^{-1/2}n^{1/4}$. As seen before (a) this growth faster than $M_n = Ch_n^{-\rho'}$.

5 Proof of the lower bound

First we show asymptotic linearity of $\hat{\vartheta}_\zeta$.

Lemma 9. *Supposing Assumptions 1 and 2 and $\zeta \in Z^{\gamma_s, \gamma_c}$ with $\gamma_s > \beta$ and $\gamma_c > (1/2 \vee \alpha) + \gamma_s$, the estimator $\hat{\vartheta}_\zeta$ with $h_n = o(n^{-1/(2\alpha+2\gamma_s)})$ is asymptotically linear with influence function $x \mapsto \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta(y)(\delta_x - \mathbb{P})(dy)$ and thus $\hat{\vartheta}_\zeta$ is Gaussian regular.*

Proof. The analysis of the bias of $\hat{\vartheta}$ in Section 4.1.1 yields

$$\begin{aligned} \hat{\vartheta} &= \vartheta + \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] \mathcal{F} K_h * \zeta(y)(\mathbb{P}_n - \mathbb{P})(dy) + o_P(n^{-1/2}) \\ &= \vartheta + \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta(y)(\mathbb{P}_n - \mathbb{P})(dy) \\ &\quad + \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] (\mathcal{F} K_h - 1) * \zeta(y)(\mathbb{P}_n - \mathbb{P})(dy) + o_P(n^{-1/2}). \end{aligned}$$

Since

$$\mathbb{E} \left[\left| \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta(d\delta_x - d\mathbb{P}) \right|^2 \right] \leq 4 \mathbb{E} \left[\int |\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta|^2 d\mathbb{P} \right]$$

is finite and $\mathbb{E}[\int \mathcal{F}^{-1}[(\varphi_\varepsilon^{-1}(-\bullet)) * \zeta](d\delta_x - d\mathbb{P})] = 0$ by (23) it suffices to show

$$\int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] (\mathcal{F} K_h - 1) * \zeta(y)(\mathbb{P}_n - \mathbb{P})(dy) = o_P(n^{-1/2}). \quad (33)$$

For convenience we write $\psi_h := \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] (\mathcal{F} K_h - 1) * \zeta$ and let $\tau > 0$. Since (Y_j) are independent and identically distributed, we obtain

$$\begin{aligned} \mathbb{P} \left(\left| n^{1/2} \int \psi_h(y)(\mathbb{P}_n - \mathbb{P})(dy) \right| > \tau \right) &\leq \tau^{-2} n \mathbb{E} \left[\left| \int \psi_h(y)(\mathbb{P}_n - \mathbb{P})(dy) \right|^2 \right] \\ &= \tau^{-2} n \mathbb{E} \left[\int \int \psi_h(y) \overline{\psi_h(z)} (\mathbb{P}_n - \mathbb{P})(dy) (\mathbb{P}_n - \mathbb{P})(dz) \right] \\ &= \tau^{-2} n^{-1} \sum_{j,k=1}^n \mathbb{E} \left[\int \int \psi_h(y) \overline{\psi_h(z)} (\delta_{Y_j} - \mathbb{P})(dy) (\delta_{Y_k} - \mathbb{P})(dz) \right] \\ &= \tau^{-1} \mathbb{E} \left[\left| \int \psi_h(y) (\delta_{Y_j} - \mathbb{P})(dy) \right|^2 \right] \\ &\leq 4\tau^{-1} \int |\psi_h(y)|^2 \mathbb{P}(dy). \end{aligned}$$

By uniform integrability of ψ_h^2 with respect to \mathbb{P} by (23) and pointwise convergence $\psi_h \rightarrow 0$ as $h \rightarrow 0$ we conclude $\int |\psi_h(y)|^2 \mathbb{P}(dy) \rightarrow 0$ and thus (33). From asymptotic linearity follows Gaussian regularity by Proposition 2.2.1 of [1]. \square

Let us now briefly discuss the consequence of Assumption 4 in terms of Fourier multipliers. Standard calculus yields $|(\varphi_\varepsilon^{-1})^{(k)}(u)| \lesssim \langle u \rangle^{\beta-k}$ for $k = 0, \dots, (\lfloor \beta \rfloor \vee M) + 1$. With the same arguments as in the proof of Lemma 5(i) we deduce that

$$(1 + iu)^{\beta+k} \varphi_\varepsilon^{(k)}(u) \text{ and } (1 + iu)^{-\beta+k} (\varphi_\varepsilon^{-1})^{(k)}(u) \quad (34)$$

are Fourier multipliers on $B_{p,q}^s(\mathbb{R})$ for all $s \in \mathbb{R}, p, q \in [1, \infty]$ and $k = 0, \dots, \lfloor \beta \rfloor \vee M$.

5.1 Information bound for smooth ζ

In this subsection we prove Theorem 2.

Step 1: To determine the solution of the maximization problem (10), we define $h := Sg = (g * f_\varepsilon) f_Y^{-1/2}$ with score operator S such that the Fisher information (9) satisfies $\langle Ig, g \rangle = \|h\|_{L^2}^2$. Therefore, we obtain $g = S^{-1}h = \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * (\sqrt{f_Y}h)$. Owing to the adjoint equation (12), $\langle g, \zeta \rangle = \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta) \sqrt{f_Y}h = \langle h, (S^{-1})^* \zeta \rangle$ holds. Ignoring all restrictions on g , the supremum is thus attained at

$$h^* := (S^{-1})^* \zeta = (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta) \sqrt{f_Y}. \quad (35)$$

Let us define $\bar{\beta} := \lfloor \beta + 1/2 \rfloor + 1$ and $r := \mathcal{F}^{-1}[(1 + iu)^{-\bar{\beta}} \varphi_\varepsilon^{-1}(u)]$. Because of Lemma 5(ii) we obtain $r \in L^1(\mathbb{R}) \cap L^2(\mathbb{R})$ and $\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(u)] = r * (\text{Id} - D)^{\bar{\beta}}$. Therefore, the condition $\int g = 0$, Fubini's theorem and the fundamental theorem of calculus, provided $(\sqrt{f_Y}h)^{(k)} \in L^1(\mathbb{R}), k = 0, \dots, \bar{\beta}$, imply

$$\begin{aligned} 0 &= \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * (\sqrt{f_Y}h) = \int r * \left(\sqrt{f_Y}h + \sum_{k=1}^{\bar{\beta}} \binom{\bar{\beta}}{k} (-1)^k (\sqrt{f_Y}h)^{(k)} \right) \\ &= \int r \left(\int \sqrt{f_Y}h + \sum_{k=1}^{\bar{\beta}} \binom{\bar{\beta}}{k} (-1)^k \int (\sqrt{f_Y}h)^{(k)} \right). \end{aligned}$$

For each $k = 1, \dots, \bar{\beta}$ the integrability of $(\sqrt{f_Y}h)^{(l)}, l = k-1, k$, yields then $\int (\sqrt{f_Y}h)^{(k)} = \lim_{x \rightarrow \infty} (\sqrt{f_Y}h)^{(k-1)}(x) - (\sqrt{f_Y}h)^{(k-1)}(-x) = 0$ and thus

$$0 = \int \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * (\sqrt{f_Y}h) = \int \sqrt{f_Y}h, \quad (36)$$

since $\int r = \mathcal{F}r(0) = 1$. Hence, we should project the solution h^* onto the L^2 -orthogonal space $\text{span}\{\sqrt{f_Y}\}^\perp$:

$$\begin{aligned} h^{**} &:= h^* - \frac{\langle h^*, \sqrt{f_Y} \rangle}{\|\sqrt{f_Y}\|_{L^2}^2} \sqrt{f_Y} \\ &= \left(\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta - \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta) f_Y \right) \sqrt{f_Y} \\ &= \left(\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta - \int \zeta f_X \right) \sqrt{f_Y}, \end{aligned} \quad (37)$$

where we used $\int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta) f_Y = \int \zeta f_X$ by (12). This leads to the candidate for the maximization of (10) given by

$$\begin{aligned} g^* &= S^{-1} h^{**} = S^{-1} (S^{-1})^* \zeta - \langle \zeta, f_X \rangle S^{-1} \sqrt{f_Y} = \mathbf{I}^{-1} \zeta - \langle \zeta, f_X \rangle f_X \\ &= \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * \left\{ \left(\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta \right) f_Y \right\} - \left(\int \zeta f_X \right) f_X \end{aligned}$$

and (12) yields $\langle g^*, \zeta \rangle = \langle \mathbf{I} g^*, g^* \rangle$ and the bound

$$\mathcal{I}_\zeta = \frac{\langle g^*, \zeta \rangle^2}{\langle \mathbf{I} g^*, g^* \rangle} = \int (\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta)^2 f_Y - \left(\int \zeta f_X \right)^2. \quad (38)$$

Inequality (11) holds then by the local version of the Hájek–Le Cam convolution theorem [1, Thm. 2.3.1]. It remains to check the conditions in (8), $(\sqrt{f_Y} h^{**})^{(k)} \in L^1(\mathbb{R})$ for $k = 0, \dots, \bar{\beta}$ and that the three-fold application of the adjoint equality is allowed. The latter will follow from $\sqrt{f_Y} h^{**}, f_Y \in H^{\beta^+}(\mathbb{R})$ for some $\beta^+ > \beta$.

Step 2: We prove now the integrability of $\sqrt{f_Y} h^{**} = \left(\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta - \int \zeta f_X \right) f_Y$ and its derivatives up to order $\bar{\beta}$ which makes the calculation (36) rigorous.

For convenience we denote

$$\kappa := \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}(-\bullet)] * \zeta = r * \left(\sum_{k=0}^{\bar{\beta}} \binom{\bar{\beta}}{k} (-1)^k \zeta^{(k)} \right).$$

Owing to Young's inequality together with $r \in L^1(\mathbb{R}) \cap L^2(\mathbb{R})$ and $\zeta^{(k)} \in L^2(\mathbb{R})$ for any $k \geq 0$, we obtain $\kappa \in C^s(\mathbb{R}) \cap H^s(\mathbb{R})$ for any $s \geq 0$. It suffices to show $f_Y^{(k)} \in L^1(\mathbb{R})$ for $k = 0, \dots, \bar{\beta}$. Note that by (34)

$$\|(\text{Id} + \text{D})^k f_\varepsilon\|_{L^1} \lesssim \|\mathcal{F}^{-1}[(1 - iu)^k \varphi_\varepsilon]\|_{B_{1,1}^0} \lesssim \|\mathcal{F}^{-1}[(1 - iu)^{k-\beta}]\|_{B_{1,1}^0}$$

is finite for $\beta > k$ since then $\mathcal{F}^{-1}[(1 - iu)^{k-\beta}] = \gamma_{\beta-k,1} \in B_{1,\infty}^{\beta-k}(\mathbb{R}) \subseteq B_{1,1}^0(\mathbb{R})$ by the proof of Lemma 5(ii). Recalling that $\beta \notin \mathbb{Z}$, we conclude iteratively $f_\varepsilon^{(k)} \in L^1(\mathbb{R})$ for $k = 0, \dots, \lfloor \beta \rfloor$. Therefore,

$$\|f_Y^{(\bar{\beta})}\|_{L^1} \leq \|f_X^{(\bar{\beta}-\lfloor \beta \rfloor)}\|_{L^1} \|f_\varepsilon^{(\lfloor \beta \rfloor)}\|_{L^1} < \infty$$

by Assumption 3 and similarly for derivatives of lower order.

Moreover, we conclude for $\bar{\beta}^- \in (\beta + 1/2, \bar{\beta})$ that

$$f_Y \in B_{1,1}^{\bar{\beta}^-}(\mathbb{R}) \subseteq B_{2,1}^{\bar{\beta}^- - 1/2}(\mathbb{R}) \subseteq H^{\beta^+}(\mathbb{R})$$

for some $\beta^+ > \beta$ by the embeddings (46) and (46). Since also $\kappa f_Y \in H^{\beta^+}(\mathbb{R})$, using $\kappa \in C^s(\mathbb{R})$ for $s > \beta$, we can apply the adjoint equality (12) in Step 1.

Step 3: We will show now $\|g^*/f_X\|_\infty < \infty$ which justifies $f_X \pm \tau g^* \geq 0$ for some choice of $\tau > 0$ small enough.

By Step 1 $g^* = \mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * (\kappa f_Y) - \langle \zeta, f_X \rangle f_X$. For the second term Assumption 3 implies $\|\langle \zeta, f_X \rangle f_X / f_X\|_\infty \leq \|\zeta\|_{L^2} \|f_X\|_\infty \|f_X\|_{L^1} < \infty$. Hence, we only need to show $\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * (\kappa f_Y) \lesssim f_X$. Using the Besov embedding (45), the Fourier multiplier property of (34) and the pointwise multiplier property of Besov spaces (48), we obtain for some $\beta^+ \in (\beta, \lfloor \beta \rfloor + 1)$

$$\|\mathcal{F}^{-1}[\varphi_\varepsilon^{-1}] * (\kappa f_Y)\|_\infty \lesssim \|\kappa f_Y\|_{B_{\infty,1}^\beta} \lesssim \|\kappa\|_{B_{\infty,1}^\beta} \|f_Y\|_{B_{\infty,1}^\beta} \lesssim \|\kappa\|_{C^s} \|f_Y\|_{C^{\beta^+}}.$$

for any $s > \beta$. In Step 2 we have seen that $\kappa \in C^s(\mathbb{R})$. Moreover,

$$\begin{aligned}
\|f_Y\|_{C^{\beta+}} &= \sum_{k=0}^{\lfloor \beta \rfloor} \|f_Y^{(k)}\|_{\infty} + \sup_{x \neq y} \left| \frac{f_Y^{(\lfloor \beta \rfloor)}(x) - f_Y^{(\lfloor \beta \rfloor)}(y)}{(x-y)^{\beta+ - \lfloor \beta \rfloor}} \right| \\
&\leq \sum_{k=0}^{\lfloor \beta \rfloor} \|f_X\|_{\infty} \|f_{\varepsilon}^{(k)}\|_{L^1} + \sup_{x \neq y} \int \frac{|f_X(x-z) - f_X(y-z)|}{|x-y|^{\beta+ - \lfloor \beta \rfloor}} f_{\varepsilon}^{(\lfloor \beta \rfloor)}(z) dz \\
&\leq \|f_X\|_{\infty} \sum_{k=0}^{\lfloor \beta \rfloor} \|f_{\varepsilon}^{(k)}\|_{L^1} + \|f_X\|_{C^{\beta+ - \lfloor \beta \rfloor}} \|f_{\varepsilon}^{(\lfloor \beta \rfloor)}\|_{L^1} < \infty,
\end{aligned} \tag{39}$$

using the Besov embedding $f_X \in W_1^2(\mathbb{R}) \subseteq B_{1,1}^{\beta+ - \lfloor \beta \rfloor + 1} \subseteq C^{\beta+ - \lfloor \beta \rfloor}$. Hence, $g^* \in L^{\infty}(\mathbb{R})$. Since f_X is a continuous, strictly positive function, we conclude that the quotient g^*/f_X is bounded on every compact subset of \mathbb{R} . Therefore, it suffices to estimate the tails. For $|x|$ large enough Assumption 3 implies, using again (34),

$$\begin{aligned}
\frac{|\mathcal{F}^{-1}[\varphi_{\varepsilon}^{-1}] * (\kappa f_Y)(x)|}{f_X(x)} &\lesssim |x^M (\mathcal{F}^{-1}[\varphi_{\varepsilon}^{-1}] * (\kappa f_Y))(x)| \\
&\leq \sum_{k=0}^M \binom{M}{k} \left| \mathcal{F}^{-1} [(\varphi_{\varepsilon}^{-1})^{(k)} \mathcal{F}[y^{M-k} \kappa f_Y]](x) \right| \\
&\lesssim \sum_{k=0}^M \|y^{M-k} \kappa f_Y\|_{B_{\infty,1}^{\beta+ - k}}.
\end{aligned}$$

Note that the above calculation shows that $\varphi_{\varepsilon}^{-1}$ is a Fourier multiplier on the weighted Besov space with weight function $\langle x \rangle^M$ [cf. 10, Def. 4.2.1/2 and Thm. 5.4.2]. Each term in the above sum can be estimated by

$$\begin{aligned}
&\|y^{M-k} \kappa\|_{C^s} \|f_Y\|_{C^{\beta+}} \\
&= \left\| \sum_{l=0}^{M-k} \binom{M-k}{l} (-1)^l \mathcal{F}^{-1} [(\varphi_{\varepsilon}^{-1})^{(l)} (-u) \mathcal{F}[(ix)^{M-k-l} \zeta]] \right\|_{C^s} \|f_Y\|_{C^{\beta+}},
\end{aligned}$$

where with abuse of notation $\beta^+ < \lfloor \beta \rfloor + 1$ is slightly larger in the last line and $s > \beta^+$. By (39) we have $f_Y \in C^{\beta^+}(\mathbb{R})$. Now, $(ix)^{M-k-l} \zeta \in \mathcal{S}(\mathbb{R})$ is again a Schwartz function and thus it suffices to show $\mathcal{F}^{-1}[(\varphi_{\varepsilon}^{-1})^{(k)} (-u) \mathcal{F} \chi] \in C^s(\mathbb{R})$ for $s > \beta$, $\chi \in \mathcal{S}(\mathbb{R})$ and $k = 0, \dots, M$. For $k = 0$ this is already done in Step 2. We proceed analogously: for any integer $s \geq 0$ we have

$$\begin{aligned}
&\|\mathcal{F}^{-1}[(\varphi_{\varepsilon}^{-1})^{(k)} (-u) \mathcal{F}[\chi]]^{(s)}\|_{\infty} \\
&= \left\| \left(\mathcal{F}^{-1} [(1+iu)^{(-\bar{\beta}+k) \wedge 0} (\varphi_{\varepsilon}^{-1})^{(k)} (-u)] * ((\text{Id} - D)^{(\bar{\beta}-k) \vee 0} \chi) \right)^{(s)} \right\|_{\infty} \\
&\leq \|(1+iu)^{(-\bar{\beta}+k) \wedge 0} (\varphi_{\varepsilon}^{-1})^{(k)} (-u)\|_{L^2} \|D^s (\text{Id} - D)^{(\bar{\beta}-k) \vee 0} \chi\|_{L^2} \\
&\lesssim \|\langle u \rangle^{((- \bar{\beta} + k) \wedge 0) + \beta - k}\|_{L^2} \|D^s (\text{Id} - D)^{(\bar{\beta}-k) \vee 0} \chi\|_{L^2}.
\end{aligned}$$

Owing to $\bar{\beta} > \beta + 1/2$, the first factor is finite since

$$((-\bar{\beta} + k) \wedge 0) + \beta - k \leq \begin{cases} \beta - \bar{\beta} < -1/2, & \text{for } \bar{\beta} \geq k, \\ \bar{\beta} - 1/2 - k < -1/2, & \text{for } \bar{\beta} < k \end{cases}$$

and the second factor is the L^2 -norm of a Schwartz function and thus finite, too.

5.2 Approximation lemma

To prove convergence of the information bounds it suffices to show that

$$\langle g_n^*, \zeta \rangle \rightarrow \langle (S^*)^{-1} \zeta, (S^*)^{-1} \zeta \rangle - \langle \zeta, f_X \rangle^2 \quad \text{and} \quad (40)$$

$$\langle Sg_n^*, Sg_n^* \rangle = \langle g_n^*, \zeta_n \rangle \rightarrow \langle (S^*)^{-1} \zeta, (S^*)^{-1} \zeta \rangle - \langle \zeta, f_X \rangle^2 \quad (41)$$

where we used the equality $\langle g_n^*, \zeta_n \rangle = \langle \mathbf{I} g_n^*, g_n^* \rangle = \langle Sg_n^*, Sg_n^* \rangle$, which holds naturally for the maximizer of the information bound \mathcal{I}_{ζ_n} . For (40) we note

$$\langle g_n^*, \zeta \rangle = \langle \mathbf{I}^{-1} \zeta_n, \zeta \rangle - \langle \zeta_n, f_X \rangle \langle \zeta, f_X \rangle = \langle (S^*)^{-1} \zeta_n, (S^*)^{-1} \zeta \rangle - \langle \zeta_n, f_X \rangle \langle \zeta, f_X \rangle$$

where the Cauchy–Schwarz inequality yields

$$|\langle f_X, \zeta_n - \zeta \rangle| = |\langle (S^*)^{-1}(\zeta_n - \zeta), \sqrt{f_Y} \rangle| \leq \| (S^*)^{-1}(\zeta_n - \zeta) \|_{L^2} \rightarrow 0 \quad (42)$$

and

$$\begin{aligned} |\langle (S^*)^{-1}(\zeta_n - \zeta), (S^*)^{-1} \zeta \rangle| &= |\langle (S^*)^{-1}(\zeta_n - \zeta), (S^*)^{-1} \zeta \rangle| \\ &\leq \| (S^*)^{-1}(\zeta_n - \zeta) \|_{L^2} \| (S^*)^{-1} \zeta \|_{L^2} \rightarrow 0 \end{aligned}$$

as $n \rightarrow \infty$. Analogously follows (41), where we use that the assumption of the lemma implies $\langle (S^*)^{-1} \zeta_n, (S^*)^{-1} \zeta_n \rangle \rightarrow \langle (S^*)^{-1} \zeta, (S^*)^{-1} \zeta \rangle$ as $n \rightarrow \infty$. The second part of the claim $\vartheta_{\zeta_n} \rightarrow \vartheta_\zeta$ has already been shown in the estimate (42).

5.3 Information bound for non-regular ζ

To prove the efficiency of $\hat{\vartheta}_t$ for $t \in \mathbb{R}$ in Theorem 4, it suffices by Lemma 3 and (35) to show

$$\langle (S^*)^{-1}(\zeta_n - \zeta), (S^*)^{-1}(\zeta_n - \zeta) \rangle^{1/2} = \| \mathcal{F}[\varphi_\varepsilon^{-1}(-\bullet)] * (\zeta_n - \zeta) \|_{L^2(\mathbb{P})} \rightarrow 0 \quad (43)$$

as $n \rightarrow \infty$. Using the moment bound (23) replacing $\mathcal{F} K_h$ by 1, we obtain

$$\| \mathcal{F}[\varphi_\varepsilon^{-1}(-\bullet)] * (\zeta_n - \zeta) \|_{L^2(\mathbb{P})} \lesssim \| \zeta_n - \zeta \|_{Z^{\beta+\delta, 1/2+\beta+\delta}}.$$

By assumption we have $Z^{\beta+\delta, 1/2+\beta+\delta} \subseteq Z^{\gamma_s, \gamma_c}$ for δ small enough. Because the space of Schwartz-functions is dense in every Sobolev space $H^s(\mathbb{R})$, $s \geq 0$, $\mathcal{S}(\mathbb{R})$ is also dense in Z^{γ_s, γ_c} and thus the information bound (11) holds for all $\zeta \in Z^{\gamma_s, \gamma_c}$. Finally, applying Theorem 25.48 of [29] and Theorem 7 from above completes the proof of Theorem 4.

A Appendix: Function spaces

Let us define the L^p -Sobolev space for $p \in (0, \infty)$ and $m \in \mathbb{N}$

$$W_p^m(\mathbb{R}) := \left\{ f \in L^p(\mathbb{R}) \mid \sum_{k=0}^m \| f_X^{(k)} \|_{L^p} < \infty \right\}$$

In particular, $W_p^0(\mathbb{R}) = L^p(\mathbb{R})$. Due to the Hilbert space structure, the case $p = 2$ is crucial. It can be described equivalently with the notation $\langle u \rangle = (1 + u^2)^{1/2}$ by, $\alpha \geq 0$,

$$H^\alpha(\mathbb{R}) := \left\{ f \in L^2(\mathbb{R}) \mid \| f \|_{H^\alpha}^2 := \int \langle u \rangle^{2\alpha} | \mathcal{F} f(u) |^2 du < \infty \right\}$$

which we call Sobolev space, too. Obviously, $W_2^m(\mathbb{R}) = H^m(\mathbb{R})$. Also frequently used are the Hölder spaces. Denoting the space of all bounded, continuous functions with values in \mathbb{R} as $C(\mathbb{R})$ we define, $\alpha \geq 0$,

$$C^\alpha(\mathbb{R}) := \left\{ f \in C(\mathbb{R}) \mid \|f\|_{C^\alpha} := \sum_{k=0}^{[\alpha]} \|f^{(k)}\|_\infty + \sup_{x \neq y} \frac{|f^{([\alpha])}(x) - f^{([\alpha])}(y)|}{|x - y|^{\alpha - [\alpha]}} < \infty \right\},$$

where $[\alpha]$ denotes the largest integer smaller or equal to α . A unifying approach which contains all function spaces defined so far, is given by Besov spaces [28, Sect. 2.3.1] which we will discuss in the sequel. Let $\mathcal{S}(\mathbb{R})$ be the Schwartz space of all rapidly decreasing infinitely differentiable functions with values in \mathbb{C} and $\mathcal{S}'(\mathbb{R})$ its dual space, that is the space of all tempered distributions. Let $0 < \psi \in \mathcal{S}(\mathbb{R})$ with $\text{supp } \psi \subseteq \{x \mid 1/2 \leq |x| \leq 2\}$ and $\psi(x) > 0$ if $\{x \mid 1/2 < |x| < 2\}$. Then define $\varphi_j(x) := \psi(2^{-j}x)(\sum_{k=-\infty}^{\infty} \psi(2^{-k}x))^{-1}$, $j = 1, 2, \dots$, and $\varphi_0(x) := 1 - \sum_{j=1}^{\infty} \varphi_j(x)$ such that the sequence $\{\varphi_j\}_{j=0}^{\infty}$ is a smooth resolution of unity. In particular, $\mathcal{F}^{-1}[\varphi_j \mathcal{F} f]$ is an entire function for all $f \in \mathcal{S}'(\mathbb{R})$. For $s \in \mathbb{R}$ and $p, q \in (0, \infty]$ the Besov spaces are defined by

$$B_{p,q}^s := \left\{ f \in \mathcal{S}'(\mathbb{R}) \mid \|f\|_{B_{p,q}^s} := \left(\sum_{j=0}^{\infty} 2^{sjq} \|\mathcal{F}^{-1}[\varphi_j \mathcal{F} f]\|_{L^p}^q \right)^{1/q} < \infty \right\}.$$

We omit the dependence of $\|\bullet\|_{B_{p,q}^s}$ to ψ since any function with the above properties defines an equivalent norm. Setting the Besov spaces in relation to the more elementary function spaces, we first note that the Schwartz functions $\mathcal{S}(\mathbb{R})$ are dense in every Besov space $B_{p,q}^s$ with $p, q < \infty$ and $H^\alpha(\mathbb{R}) = B_{2,2}^\alpha(\mathbb{R})$ as well as $C^\alpha(\mathbb{R}) = B_{\infty,\infty}^\alpha(\mathbb{R})$, where the latter holds only if α is not an integer [28, Thms. 2.3.3 and 2.5.7]. Frequently used are the following continuous embeddings which can be found in [28, Sect. 2.5.7, Thms. 2.3.2(1), 2.7.1]: For $p \geq 1, m \in \mathbb{Z}$

$$B_{p,1}^m(\mathbb{R}) \subseteq W_p^m(\mathbb{R}) \subseteq B_{p,\infty}^m(\mathbb{R}) \quad \text{and} \quad B_{\infty,1}^0(\mathbb{R}) \subseteq L^\infty(\mathbb{R}) \subseteq B_{\infty,\infty}^0(\mathbb{R}) \quad (44)$$

and for $s \geq 0$

$$B_{\infty,1}^s(\mathbb{R}) \subseteq C^s(\mathbb{R}) \subseteq B_{\infty,\infty}^s(\mathbb{R}). \quad (45)$$

Furthermore, for $0 < p_0 \leq p_1 \leq \infty, q \geq 0$ and $-\infty < s_1 \leq s_0 < \infty$

$$B_{p_0,q}^{s_0}(\mathbb{R}) \subseteq B_{p_1,q}^{s_1}(\mathbb{R}) \quad \text{if} \quad s_0 - \frac{1}{p_0} \geq s_1 - \frac{1}{p_1} \quad (46)$$

and for $0 < p, q_0, q_1 \leq \infty$ and $-\infty < s_1 < s_0 < \infty$

$$B_{p,q_0}^{s_0}(\mathbb{R}) \subseteq B_{p,q_1}^{s_1}(\mathbb{R}). \quad (47)$$

Another important relation is the pointwise multiplier property of Besov spaces [28, (24) on p. 143] that is

$$\|fg\|_{B_{p,q}^s} \lesssim \|f\|_{B_{\infty,q}^s} \|g\|_{B_{p,q}^s} \quad (48)$$

for $s > 0, 1 \leq p \leq \infty$ and $0 < q \leq \infty$.

The Besov norm of a convolution can be bounded by Lemma 7 (i) in [25]. Let $1 \leq p, q, r, s \leq \infty$, $-\infty < \alpha, \beta < \infty$, $0 \leq 1/u = 1/p + 1/r - 1 \leq 1$, $0 \leq 1/v = 1/q + 1/s \leq 1$. For $f \in B_{p,q}^\alpha(\mathbb{R})$ and $g \in B_{r,s}^\beta(\mathbb{R})$

$$\|f * g\|_{B_{u,v}^{\alpha+\beta}} \lesssim \|f\|_{B_{p,q}^\alpha} \|g\|_{B_{r,s}^\beta}. \quad (49)$$

Using for any function $f : \mathbb{R} \rightarrow \mathbb{R}$ and $h \in \mathbb{R}$ the difference operators $\Delta_h^1 f(x) := f(x+h) - f(x)$ and $(\Delta_h^l f)(x) := \Delta_h^1(\Delta_h^{l-1} f)(x)$, $l \in \mathbb{N}$, the Besov can be equivalently described by

$$\|f\|_{B_{pq}^s} \sim \|f\|_{L^p} + \|f\|_{\dot{B}_{pq}^s} \quad \text{with} \quad \|f\|_{\dot{B}_{pq}^s} := \left(\int |h|^{-sq-1} \|\Delta_h^M f\|_{L^p}^q dh \right)^{1/q}$$

for $s > 0, p, q \geq 1$ and any integer $M > s$ [28, Thm. 2.5.12]. The space of all $f \in \mathcal{S}'(\mathbb{R})$ for which $\|f\|_{\dot{B}_{pq}^s}$ is finite is called homogeneous Besov space $\dot{B}_{pq}^s(\mathbb{R})$ [28, Def. 5.1.3/2, Thm. 2.2.3/2] and thus $B_{pq}^s = L^p(\mathbb{R}) \cap \dot{B}_{pq}^s(\mathbb{R})$ for $s > 0, p, q \geq 1$. Of interest is the relation of homogeneous Besov spaces to functions of bounded p -variation. Let $\mathcal{BV}_p(\mathbb{R})$ denote the space of measurable functions $f : \mathbb{R} \rightarrow \mathbb{R}$ such that there is a function g which coincides with f almost everywhere and satisfies

$$\sup \left\{ \sum_{i=1}^n |g(x_i) - g(x_{i-1})|^p \mid -\infty < x_1 < \dots < x_n < \infty, n \in \mathbb{N} \right\} < \infty$$

and we define $BV_p(\mathbb{R})$ as the quotient set $\mathcal{BV}_p(\mathbb{R})$ modulo equality almost everywhere. Then,

$$\dot{B}_{p1}^{1/p}(\mathbb{R}) \subseteq BV_p(\mathbb{R}) \subseteq \dot{B}_{p,\infty}^{1/p}(\mathbb{R}), \quad \text{for } p > 1 \quad (50)$$

by [4, Thm. 5]. For $p = 1$ holds by [14, Lem. 8]

$$BV_1(\mathbb{R}) \cap L^1(\mathbb{R}) \subseteq B_{1,\infty}^1(\mathbb{R}). \quad (51)$$

References

- [1] Bickel, P. J., C. A. Klaassen, Y. Ritov, and J. A. Wellner (1998). *Efficient and Adaptive Estimation for Semiparametric Models*. Springer, New York.
- [2] Bickel, P. J. and Y. Ritov (2003). Nonparametric estimators which can be “plugged-in”. *Ann. Statist.* 31(4), 1033–1053.
- [3] Bissantz, N., L. Dümbgen, H. Holzmann, and A. Munk (2007). Non-parametric confidence bands in deconvolution density estimation. *J. R. Stat. Soc. Ser. B Stat. Methodol.* 69(3), 483–506.
- [4] Bourdaud, G., M. Lanza de Cristoforis, and W. Sickel (2006). Superposition operators and functions of bounded p -variation. *Rev. Mat. Iberoam.* 22(2), 455–487.
- [5] Butucea, C. and F. Comte (2009). Adaptive estimation of linear functionals in the convolution model and applications. *Bernoulli* 15(1), 69–98.
- [6] Dattner, I., A. Goldenshluger, and A. Juditsky (2011). On deconvolution of distribution functions. *The Annals of Statistics* 39(5), 2477–2501.

- [7] de la Peña, V. H. and E. Giné (1999). *Decoupling: From Dependence to Independence*. Springer, New York.
- [8] Dudley, R. M. (1992). Fréchet differentiability, p -variation and uniform Donsker classes. *Ann. Probab.* 20(4), 1968–1982.
- [9] Dudley, R. M. (1999). *Uniform Central Limit Theorems*. Cambridge University Press, Cambridge.
- [10] Edmunds, D. E. and H. Triebel (1996). *Function Spaces, Entropy Numbers, Differential Operators*. Cambridge University Press, Cambridge.
- [11] Fan, J. (1991a). Asymptotic normality for deconvolution kernel density estimators. *Sankhyā: The Indian Journal of Statistics* 53(1), 97–110.
- [12] Fan, J. (1991b). On the Optimal Rates of Convergence for Nonparametric Deconvolution Problems. *The Annals of Statistics* 19(3), 1257–1272.
- [13] Föllmer, H. and A. Schied (2004). *Stochastic finance: an introduction in discrete time* (extended ed.), Volume 27 of *de Gruyter Studies in Mathematics*. Walter de Gruyter & Co., Berlin.
- [14] Giné, E. and R. Nickl (2008). Uniform central limit theorems for kernel density estimators. *Probab. Theory Related Fields* 141(3-4), 333–387.
- [15] Giné, E. and R. Nickl (2009). Uniform limit theorems for wavelet density estimators. *Ann. Probab.* 37(4), 1605–1646.
- [16] Girardi, M. and L. Weis (2003). Operator-valued Fourier multiplier theorems on Besov spaces. *Mathematische Nachrichten* 251(1), 34–51.
- [17] Goldenshluger, A. and S. V. Pereverzev (2003). On adaptive inverse estimation of linear functionals in Hilbert scales. *Bernoulli* 9(5), 783–807.
- [18] Hall, P. and S. N. Lahiri (2008). Estimation of distributions, moments and quantiles in deconvolution problems. *Ann. Statist.* 36(5), 2110–2134.
- [19] Klenke, A. (2007). *Probability Theory: A Comprehensive Course (Universitext)* (1st ed.). Springer, London.
- [20] Kosorok, M. R. (2008). *Introduction to Empirical Processes and Semiparametric Inference* (1st ed.). Springer Series in Statistics. Springer, New York.
- [21] Lounici, K. and R. Nickl (2011). Global uniform risk bounds for wavelet deconvolution estimators. *Ann. Statist.* 39(1), 201–231.
- [22] Nickl, R. (2006). Empirical and Gaussian processes on Besov classes. In *High dimensional probability*, Volume 51 of *IMS Lecture Notes Monogr. Ser.*, pp. 185–195.
- [23] Nickl, R. and B. M. Pötscher (2007). Bracketing metric entropy rates and empirical central limit theorems for function classes of Besov- and Sobolev-type. *J. Theoret. Probab.* 20(2), 177–199.

- [24] Nickl, R. and M. Reiß (2012). A Donsker Theorem for Lévy Measures. arXiv 1201.0590.
- [25] Qui, B. H. (1981). Bernstein’s theorem and translation invariant operators. *Hiroshima Math. J.* 11(1), 81–96.
- [26] Radulović, D. and M. Wegkamp (2000). Weak convergence of smoothed empirical processes: beyond Donsker classes. In *High dimensional probability, II*, Volume 47 of *Progr. Probab.*, pp. 89–105. Birkhäuser, Boston.
- [27] Schmidt–Hieber, J., A. Munk, and L. Dümbgen (2012). Multiscale Methods for Shape Constraints in Deconvolution: Confidence Statements for Qualitative Features. arXiv 1107.1404.
- [28] Triebel, H. (2010). *Theory of Function Spaces*. Birkhäuser Verlag, Basel. Reprint of the 1983 Edition.
- [29] van der Vaart, A. W. (1998). *Asymptotic statistics*. Cambridge Series in Statistical and Probabilistic Mathematics. Cambridge University Press.
- [30] van der Vaart, A. W. and J. A. Wellner (1996). *Weak convergence and empirical processes*. Springer Series in Statistics. Springer, New York.
- [31] van Es, B. and H.-W. Uh (2005). Asymptotic normality of kernel-type deconvolution estimators. *Scand. J. Statist.* 32(3), 467–483.

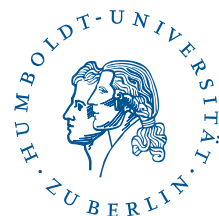
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