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Working Paper
Trends in U.S. Family Income Mobility, 1969-2006

Working Papers, No. 11-10

Provided in Cooperation with:
Federal Reserve Bank of Boston

Suggested Citation: Bradbury, Katharine (2011) : Trends in U.S. Family Income Mobility, 1969-2006, Working Papers, No. 11-10, Federal Reserve Bank of Boston, Boston, MA

This Version is available at:
http://hdl.handle.net/10419/55599

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Trends in U.S. Family Income Mobility, 1969–2006
Katharine Bradbury

Abstract:
Much of America’s promise is predicated on economic mobility—the idea that people are not limited or defined by where they start, but can move up the economic ladder based on their efforts and accomplishments. Family income mobility—changes in individual families’ income positions over time—is one indicator of the degree to which the eventual economic wellbeing of any family is tethered to its starting point. In the United States, family income inequality has risen from year to year since the mid-1970s; given this rising cross-sectional inequality, changes over time in mobility determine the degree to which long-term income is also increasingly unequally distributed.

Using data from the Panel Study of Income Dynamics and a number of mobility concepts and measures drawn from the literature, this paper examines family income mobility levels and trends for U.S. working-age family heads and spouses during the time span 1969–2006, based on a post-tax, post-transfer concept of income adjusted for family size. By most measures, mobility is lower in more recent periods (1995–2005) than in the late seventies and the eighties (the 1977–1987 or 1981–1991 periods). Comparing results based on pre-government income suggests that an increasingly redistributive tax and transfer system contributed to rising mobility into the 1980s, but that its impact has since waned. Overall, the evidence indicates that over the 1969-to-2006 time span, family income mobility across the distribution decreased, families’ later-year incomes increasingly depended on their starting place, and the distribution of families’ lifetime incomes became less equal.

JEL codes: D31, D63, I32, J15
Keywords: Income mobility, income inequality, income distribution

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This paper presents preliminary analysis and results intended to stimulate discussion and critical comment. The views expressed herein are those of the author and do not indicate concurrence by other members of the research staff or principals of the Federal Reserve Bank of Boston or the Federal Reserve System.

This paper, which may be revised, is available on the web site of the Federal Reserve Bank of Boston at http://www.bos.frb.org/economic/wp/index.htm.

This paper builds heavily on earlier joint work with Jane Katz, Officer and Director of Education Programs at the Federal Reserve Bank of New York; while she is not co-author, her insights added immeasurably to the research. Peter Gottschalk generously provided edited data files from the Panel Study of Income Dynamics, without which this analysis would have been much more difficult to complete; Dean Lillard helped me understand the CNEF files at Cornell University. I am also grateful to Mary Burke, Jane Little, Bob Triest, Dan Aaronson, two anonymous referees, participants at the May 2009 conference of the System Committee on Applied Microeconomics and at the November 2009 APPAM meetings for helpful comments on earlier drafts, to participants in a Boston Fed departmental seminar for comments on a very early presentation of this work, and to Jessamyn Fleming, Charu Gupta, and Ryan Kessler for outstanding research assistance and comments. I am responsible for any remaining errors.

This version: October 20, 2011
Much of America’s promise is predicated on the existence of economic mobility—the idea that people are not limited or defined by their current circumstances, but can move up the income ladder based on their effort and accomplishments. Can a poor individual or family rise into the middle class, or is longer-term economic status limited by one’s economic situation at a given point in time?

Changes in economic mobility are of particular consequence when economic disparities among families are increasing over time, as has been the case in the United States in recent decades. If family income inequality is increasing, changes in the degree to which families move up and down can either offset or amplify longer-term inequality—and loosen or tighten the link between a family’s circumstances in any given year and its later outcomes. Other things being equal, an economy with rising mobility—one in which families move increasingly frequently or traverse increasingly greater distances up and down the income ladder—will result in a more equal distribution of lifetime incomes than an economy with declining mobility.

This paper examines time patterns of family income mobility for U.S. working-age family heads and spouses between 1969 and 2006, using data from the Panel Study of Income Dynamics (PSID) and a number of mobility concepts and measures, including a measure of the degree to which mobility equalizes long-term incomes. Calculating these measures for overlapping 10-year periods, the paper documents mobility levels and trends based on pre- and post-government measures of family income adjusted for family size.

By and large, different mobility measures and income measures yield similar pictures of mobility trends. Most mobility measures indicate that family income mobility was lower in more recent periods (the 1990s into the early 2000s) than in earlier periods. Comparing 1977–1987 or 1981–1991, when many of the measures peaked, with the most recent period (1995–2005), mobility declined to a statistically significant degree according to most measures, both overall and for individuals beginning near the top or bottom of the income distribution. However, comparing periods ending in the most recent 10 or so years (1985–1995 through 1995–2005), depending on the mobility measure employed, trends are difficult to discern, except that mobility was lower in the last period (1995–2005 or 1996–2006) than in the period immediately
before (1993–2003 or 1994–2004). Comparing these post-government trends with results based on pre-government income suggests that the redistributive impact of the U.S. tax and transfer system may have contributed more positively to mobility in the 1970s than it has since 1981.

As the opening paragraphs above indicate, a key reason for studying mobility trends is a concern with lifetime income inequality. Although the calculations reported here track individuals over 10 years (far less than a lifetime), the distribution of long-term income (family income averaged over 10 years, or even 16 years) was much more unequal in recent periods than in the 1970s. The trend has been visibly steep, regardless of the inequality measure or income measure employed. Thus, whether the declines in measured mobility are significantly different from zero or not, mobility levels and trends have not been sufficient to offset the considerable rise in short-term inequality.

The paper proceeds as follows: The first section below provides an overview of existing literature related to changes over time in family income mobility. The following section discusses concepts of income mobility and outlines the concepts and measures used in this paper (an appendix provides actual definitions of the measures). The next section describes the data and sample used in the analysis: the “base case” measure of income, how it is adjusted for family size, and the mobility “periods” examined. Three sections follow that report and discuss results according to various mobility measures and examine their time trends. The next two sections introduce comparisons of the post-government base case with the pre-government income case, looking at “mobility” for individuals from their pre-government income status to their post-government income status within a year and at mobility measured over 10-year periods based on pre-government income. The final section summarizes and discusses the results.

**Literature Review**

An extensive literature explores the extent of economic mobility in the United States and other nations. Much of that research focuses on earnings mobility, examining how individuals’ earnings change over time. Another set of papers focuses on family income mobility, measuring the degree to which individuals’ family incomes change from one point in time to another, and
a subset of that research investigates how family income mobility patterns have changed over time. The current paper contributes to this latter literature, measuring U.S. family income mobility with a variety of measures and in many overlapping time periods in order to explore their time patterns: It asks whether U.S. family heads and spouses were more or less likely to move up and down the family income ladder in recent periods than in earlier decades, comparing, for example, mobility during the 1996–2006 period (most recent available) with that during 1969–1979 (earliest available) or with some other 10-year period in the middle of the 1969–2006 span. And it explores whether the answer to this question depends on the specific measures of mobility or of family income or on a focus on specific parts of the income distribution.


\(^{1}\) Fields and Ok (1999) compare the early part of two decades, but use only absolute measures—income flux and directional income change—based on individual changes in log income (which they describe as “a particular facet of the multi-faceted notion of income mobility,” p. 467). They say their finding of a statistically significant increase in (family) income flux in the United States between the 1969–1976 and 1979–1986 periods is complementary to “earlier findings of others who demonstrated that relative mobility in the United States has been unchanged or falling over the same period of time” (p. 457).
Looking at even earlier periods, Sawhill and Condon (1992), Hungerford (1993), and Gittleman and Joyce (1999) report little change in overall family income mobility between the 1970s and the 1980s. The latter two studies, however, find “subtle differences” (Hungerford’s term), suggesting that some groups were less upwardly mobile in the later decade. Focusing on young children, Gottschalk and Danziger (2001) also find little difference in family income mobility between the 1970s and the 1980s, although children whose families began with low income (poorest two quintiles) or high income (richest quintile) were somewhat less mobile in the later period.

The results in this paper are broadly consistent with this literature, and the “Results” and “Time Trends” sections below provide more detailed comparisons. The paper reports a variety of measures on a consistent basis for overlapping 10-year time periods over a 37-year time span and is able to compare and contrast the trends in the literature with those observed here. Thus, this paper contributes to the literature by examining mobility over a considerably longer time span and investigating a broader range of mobility measures, and it offers a new categorization of mobility measures as well as comparisons of mobility based on pre- and post-government incomes.

**Mobility Concepts and Measures**

In broad terms, mobility is the pace and degree to which individuals’ or families’ incomes (or other measures of wellbeing) change over time relative to one another or relative to the overall income distribution. Some researchers define mobility even more broadly to encompass movements that have no “relative” aspect, such as the average change in (log) incomes or the average absolute value of (log) income change. While overall income growth obviously contributes to average wellbeing, this paper treats it as a feature of income change that is distinct from mobility. Thus, the mobility measures examined in this paper all have a relative aspect, comparing families’ income changes with one another, not simply computing average changes.2

2 Fields (2008a) notes that a “difference of view—whether ‘income mobility’ includes the growth aspect of distributional change or whether ‘mobility’ is what remains after growth has been taken out—underlies
Among concepts and measures with a relative aspect, this paper emphasizes the features that help to address the questions posed at the outset: How limited is a person by his/her current circumstances and how are income changes contributing (or not) to the inequality of lifetime incomes? The degree to which end-of-period income (or position) is independent of beginning-of-period income or position is related to the idea of equal opportunity. In their “introduction to the literature” on income mobility measurement, Fields and Ok (2001) note that “origin independence seems to capture our intuitions about ‘equality of opportunity’… Viewed in this way, income mobility is a desirable notion that helps attenuate the unequal distribution of initial endowments.” (p. 561). Mobility as an equalizer of longer-term or lifetime incomes, by contrast, associates more equal long-term outcomes (not opportunities) with improved mobility.

The paper uses a variety of relative measures in order to understand whether they all tell a consistent story and whether findings from previous studies are artifacts of particular measures. Reflecting an interest in mobility during a working life, the paper examines mobility over the long term (10-year periods); it does not address shorter-term “volatility”—shocks to family incomes from year to year—nor even longer-term intergenerational mobility—how much a person’s prime-age family income level (or position) depends on the corresponding level (position) of his/her parents during his/her childhood.

The sections that follow discuss two ways of categorizing mobility concepts. When cross-classified, the two category types yield the two-by-two matrix shown in Table 1, which lists the specific measures from the literature used in this paper; the appendix defines the measures.

much of the mobility literature, but rarely is made explicit.” For this reason, I am being explicit that the definition used in this paper excludes that pure average-growth aspect. Van Kerm (2002) explicates the distinction as follows: he says the Fields and Ok (1999) flux concept of mobility, depending as it does on absolute change in log income, “differs from most approaches to income mobility measurement in one important respect: income mobility is seen as the juxtaposition of isolated individual experiences and not as an intrinsically social phenomenon where it is individual experiences relative to the experiences of others that matter.” (p. 232, emphasis added). In these terms, the mobility investigated in this paper, like “most approaches,” is a social phenomenon.

3 “Lifetime” income typically refers to a person’s average income over a number of years; the current analysis focuses on average income in a 10-year period during the working-age years.
Position-relative and dollar-relative mobility

One categorization is built around the specific aspects of income movement considered; these define the two horizontal panels in Table 1.

Position-relative mobility (upper panel) refers to individuals moving among positions in the distribution of family incomes between the beginning and end of a period, where positions are defined in a purely relative way. Position-relative measures are typically based on transition matrices showing position in terms of quintile, decile, or other quantile at the beginning and the end of a period, but they can also be expressed more granularly in terms of rank. What makes them position-relative is that they do not reflect the dollar size of incomes or the spread of the income distribution, but only rank.

Dollar-relative mobility (lower panel) refers to changes in individuals’ family incomes and associated changes in their relative positions. This concept encompasses income movement both relative to some real standard of wellbeing or purchasing power and relative to the (possibly changing) overall distribution of income. Dollar-relative mobility can reflect changes in the structure of rewards in the economy—the addition, for example, of much higher places at the top of the distribution—as well as changes in individuals’ access to these rewards.

Dollar-relative measures are useful for summarizing how much movement the average family experiences—both relative to other families and in terms of absolute (dollar) income change—taking into account changes in average income levels and the inequality of the income distribution. Some of these measures directly indicate the degree to which longer-term incomes are equalized by income movements in a period; others shed light on the degree to which an individual’s initial family income position persists.

Overall and origin-specific mobility

The second categorization is based on the individuals for whom family income movements are considered—part or all of the initial distribution; these define the two vertical panels in Table 1.

Overall mobility (left-hand column) summarizes the transition process from the economy-wide set of individual observations of wellbeing—whether income or rank—at the start of a
period to the corresponding set at the end of that period. Overall mobility measures attempt to quantify the extent to which the entire economy is characterized by families with persistent positions versus families experiencing widespread movement. They also may summarize the economy-wide association between families’ changes in rank and absolute income or the degree to which families’ patterns of movement are making longer-term incomes more or less equal.

*Origin-specific mobility* (right-hand column) summarizes movements of subgroups defined by their position (rank, income) in the distribution at the beginning of the period. Origin-specific mobility measures attempt to quantify the extent to which those who start at the bottom/top (or, more generally, in any specific part of the distribution) move up/down either relatively (rank, position) or absolutely (by crossing some real dollar threshold) by the end of the period.

Origin-specific measures are of interest for several reasons, including concerns about the ability of the poor to escape the bottom rungs of the income ladder and concerns about stability at the top, as evidence of unequal opportunity or a lack of meritocracy. Along these lines, Burkhauser and Couch (2009) note that studies of origin-specific mobility “have been motivated by an interest in whether those at the top and bottom of society are differentially mobile but also by an interest in the degree to which spells in poverty impact future movement up the income distribution.” (p. 523). In this regard, a recent Pew Trusts survey found that "a majority of Americans believe that the lack of upward mobility from the bottom rung of the income ladder is a major problem for this country, while they are relatively unconcerned about how little downward mobility there is from the top" (Economic Mobility Project 2009, p. 4).

**Data and Sample**

The mobility measures discussed above and defined in the appendix are calculated using data from the Panel Study of Income Dynamics (PSID), which has collected information on the incomes and characteristics of individuals and their families since 1968. The survey was conducted every year from 1968 through 1997 and every other year thereafter; the most recent survey for which income data are available, conducted in 2007, provides data on incomes in 2006.
Each period’s measures include all individuals who meet the following criteria at both the beginning and end of the period (or throughout the period for measures that use all years):

- individual is a family head or spouse
- both the family’s head and spouse (if present) are between 16 and 62 years old
- family income data are not missing.

In addition, at the beginning of a period, the individual’s family may not be a “split-off,” meaning that the family must have been separate from the head’s parents’ family for at least a year. Thus, the sample (approximately 2,500 to 4,000 observations) changes for each period. The observations are weighted using individual weights, to correct for the PSID’s over-sampling of the bottom of the income distribution.

The paper examines two measures of family income. One is pre-government family income, which comprises pretax, pre-transfer money income of all family members combined, including wages, salaries, rent, interest, dividends, farm and business income, pensions, alimony, child support, and help from relatives and others. The second income measure is post-government family income, which subtracts net income taxes (“net” meaning that it also reflects the potentially positive addition of the earned income tax credit for low-income families with at least one worker) and adds in the cash value of food stamps along with social security benefits and all government cash transfers, such as welfare. The pre-government and post-government measures are drawn from Cornell University’s Cross-National Equivalent File for the PSID.

Income and other dollar-based measures are adjusted for inflation using the CPI-U-RS.

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4 For mobility measures based on two-year average endpoints (the base case in Figures 1 through 4), heads and spouses are between 16 and 62 years old in both years used in computing both endpoints. Thus, for example, the measures shown in Figure 1 for the 1995–2005 period are based on heads and spouses who are 16 to 62 years of age in 1994, 1996, 2004, and 2006; hence they are 16 to 50 years of age in 1994 and 28 to 62 years old in 2006.

5 Since income is reported for the calendar year prior to the survey, split-off families’ reported income includes the full- or part-year income of the parental family. The exclusion of split-off families thus insures that measured mobility reflects the individual’s independent family income changes, not the income movements that result when children move from parental to independent family income.

6 After exclusion of observations with zero weight in either year, the individual weight is averaged over all the years included in each period’s measure.

7 Cornell University imputes net taxes via the TAXSIM model, and assumes that all taxpaying units take the standard deduction and all eligible families file for the earned income tax credit.
For post-government income, the observations in the highest and lowest (weighted) percentile of each year’s head-and-spouse distribution are trimmed out. This eliminates any top-coded and bottom-coded observations as well as the most extreme measurement errors. Only the highest 1 percent are trimmed from the pre-government income measure because it is impossible to trim the poorest percentile without trimming more than 1 percent, and because the income values at the bottom have a reasonable interpretation. From 1½ to 5 percent (depending on year) of weighted heads and spouses have zero pre-government income, indicating that the family has no income in the absence of government transfers or social security. The zeros therefore do not represent extreme values or likely measurement error, and dropping all of them would leave a biased sample of pre-government incomes. When comparing measures based on pre- and post-government income, the measures incorporate only observations present in both samples.

Family income is adjusted for family composition to yield a more accurate indicator of wellbeing; for example, $45,000 represents a very different standard of living for a family of five than for a two-person family. The adjustment divides family income by the square root of family size, an equivalence scale used in a number of research papers.

In most of the analysis that follows, the reported measures relate to a “base case,” in which the following holds:

- The measure of income is post-government income adjusted for family composition using the square root of family size;

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8 After 1996, some observations have negative pre-government income; these are set to zero for consistency with the data for 1969 through 1996 (when the PSID set all negative income values to zero or $1).

9 Burkhauser et al. (1996) explore the implications of using several possible equivalence scales in measuring inequality, including the poverty line and some based on expenditures; they label the square root of family size the “International Experts” scale. Karoly and Burtless (1995) adopt the square root of family size in their exploration of Gini inequality and report only modest differences in the trend from using 0 (no adjustment of total family income) vs. 0.5 (their Appendix Table 1). Burkhauser, Larrimore, and Simon (2011) say that dividing by the square root of the income-sharing-unit’s size is “the customary procedure in the income inequality literature” (p. 13). The most widely used alternatives are to adjust family income by dividing by the Census poverty line, which varies with family size (and, over time, with inflation), or by dividing by the PSID “needs” measure, a similar scale based on the USDA “low cost” food standard.
• The period for measuring mobility is 10 years: and
• Adjusted income is averaged for two years at both the start and the end of the 10-year period.

Averaging two years of adjusted income at the period endpoints is intended to smooth some of the transitory income changes that occur on a year-to-year basis as well as reduce the effects of measurement error in single-year income. Because the data are collected only every other year after the 1996 income year (1997 survey), the two-year endpoints are calculated by averaging non-adjacent years \((t-1\) and \(t+1\)); for consistency, this approach is also used before 1996. Labels refer to year \(t\); thus, for example, “2005” is the average of 2004 and 2006 income data, so the period 1995–2005 reflects income changes between the average of 1994 and 1996 and the average of 2004 and 2006.

One set of measures—Shorrocks M and the inequality of long-term incomes on which it is based—is computed using the average of all the observations on each individual’s adjusted family income within a 10-year period rather than just the beginning and end years. Because data are available only every other year, the average includes income over the six years comprising every other year within each 10-year period (the computations are the same—averaging every other year—even in the earlier periods during which data exist for all years). For example, the most recent 10-year average measure, for 1996–2006, averages family income for each head or spouse observed in 1996, 1998, 2000, 2002, 2004, and 2006.

Results for Position-Relative Mobility

Position-relative mobility measures consider families’ movements in relative position from the beginning to the end of a period, where relative position is indicated by quintile, decile, or some other rank-based indicator. As noted earlier (and as categorized in Table 1), position-relative mobility can be measured for all individuals combined (overall) or for subsets defined by where in the distribution they begin the period (origin-specific).

Overall position-relative mobility

Figure 1 displays five position-relative overall mobility measures for the base case and
Table 2 reports the values. The time patterns of the individual measures are very similar: most of them increase somewhat or move sideways between 1970–1980 and 1977–1987 or thereabouts, decline non-monotonically through 1989–1999, level out or rise slightly, and decline again between 1991–2001 or 1993–2003 and 1995–2005. They all indicate less family income mobility for U.S. individuals in the (most recent) 1995–2005 period than in the 1970s. For example, 78 percent of individuals moved up or down one decile or more in the first few periods, while 75 to 76 percent changed deciles during the last few periods. The number of deciles the average person moved declined from 1.9 to 1.7. M-hat decreased from around 0.76 to 0.72, and the correlation of an individual’s beginning-of-period and end-of-period income rank rose from 0.62 to 0.65, indicating mobility (as measured by the inverse rank correlation) declined. The inverse adjusted contingency coefficient C* fell from 0.36 to 0.33. While these are all small changes, they are consistently negative across the position-relative overall measures.

Auten and Gee’s (2009) tax-return-based mobility matrices indicate a slight decline in M-hat (from 0.702 to 0.695) between 1987–1996 and 1996–2005; our data also show declines in these measures between 1987–1997 and 1995–2005. The quintile mobility matrices reported in Tables 7–9 of Carroll, Joulfaian, and Rider (2006) indicate that M-hat fell slightly (from 0.70 to 0.68) between the 1979–1986 and 1987–1995 periods; over this time span, Table 2 below shows declines in M-hat (for 10-year mobility) as well. For even earlier decades, Gittleman and Joyce (1999) document a gradual increase in mobility between 1967–1977 and 1981–1991; they display (graphically) the quintile immobility ratio (fraction staying in origin quintile), which moves from about 0.37 for 1967–1977 to 0.355 in 1981–1991; they also report a slight increase in mobility between two adjacent 12-year periods, 1967–1979 and 1979–1991. As noted above, Table 2 and Figure 1 below also show increases in mobility across 10-year periods through the

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10 Two of the measures have been re-scaled in the figure so that all of them can be displayed together.
11 Their National Tax Journal paper does not include mobility matrices defined over the panel population for both periods, but their working paper (2007) does (Appendix Table A.5).
12 In the “time trends” section below, I compare my results with those of Acs and Zimmerman (2008) and other researchers who use adjacent decade pairs.
13 Gittleman and Joyce use PSID income adjusted by the poverty line and follow individual heads and spouses ages 25 to 64. The immobility ratio figures are based on their smoothed (“predicted”) series in the lower panel of Figure 1; their Table 5 reports quintile mobility matrices for 1967–1979 and 1979–1991.

**Origin-specific position-relative mobility**

Origin-specific mobility measures focus on the movement of those who begin in a specific segment of the initial distribution. Figure 2 shows the percentages falling out of the top quintile and rising from the bottom quintile, respectively, and also the percentages of those who start in the richest or poorest quintile and move “far” relative to the distribution, defined as beyond the adjacent quintile; the right-hand panel of Table 2 reports the measures. The first thing to note in Figure 2 is that those who begin in the poorest quintile are less likely to move up than those who begin in the richest quintile are to move down; this lower mobility of the poor applies both to any movement out of origin quintile and to far moves.

Regarding time patterns, like the overall position-relative measures, these origin-specific measures rise slightly or move sideways initially, decline gradually until 1987–1997 (poor) or 1989–1999 (rich), rise a bit, and then fall again between 1993–2003 and 1995–2005. The decline in mobility is somewhat more pronounced for those who begin rich than for those who start a period in the poorest quintile, but not so for bigger moves.

Auten and Gee’s (2007) tax-return-based mobility matrices show decreases in the percentage of the richest quintile who move down (from 41.6 to 39.6) and who move far (from 18.1 to 16.5) between 1987–1996 and 1996–2005; Table 2 below also shows declines when comparing 1987–1997 and 1995–2005. For earlier periods, Hungerford (1993) reports his data in terms of decile transition matrices, but when combined into quintiles his data show 43.6 percent of the poorest quintile moving up between 1969 and 1976, and 42.4 percent moving up between 1979 and 1986.

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14 For Gottschalk and Danziger, M-hat calculated with their quintile mobility matrices (Table 5.1) is 0.701 in 1971–1981 and 0.698 in 1981–1991; their measures use three-year average endpoints. For Sawhill and Condon, M-hat calculated for their quintile mobility matrices (Table 1) is 0.756 for 1967–1976 and 0.759 for 1977–1986.
Gottschalk and Danziger (2001) report slight declines in the fraction of children starting in the poorest or richest quintiles who move out between the 1971–1981 and 1981–1991 periods, and more noticeable declines for “far” moves by initially poorest-quintile children (falling from 14.8 to 12.4 percent) and for children starting in the richest quintile (from 18.5 to 11.3 percent). By contrast, Sawhill and Condon (1992) see higher fractions moving out of the bottom and top quintiles in 1977–1986 than in 1967–1976. Indeed, across the early periods, Figure 2 and Table 2 below also fail to show declines in origin-specific relative mobility.

Results for Dollar-Relative Mobility

*Overall dollar-relative mobility, part 1*

Figure 3 displays three measures of dollar-relative overall mobility for the base case (also see Table 3). The Gini mobility measure, which reflects the association of changes in rank and income during a period, exhibits a time pattern similar to those of the overall position-relative (that is, rank-based) measures in Figure 1, moving generally downward after being level across the first several periods.

The other two measures declined markedly between 1970–1980 and 1980–1990, rose gradually from 1980–1990 to 1991–2001 or 1993–2003, and then fell again to the last (1995–2005) period. The elasticity of end-year income with respect to beginning-year income rose from around 0.65 in the first few periods to 0.75 in the last, indicating a decline in mobility in terms of the degree to which end-year income moves independently of beginning-year income. The Fields directional mobility-as-equalizer measure also fell.\footnote{\textsuperscript{15,16}}

Figure 4 displays absolute quintile mobility measures (also Table 3). Absolute quintile

\footnote{15 Note that the Fields measure—which, being directional, can be negative—is positive but close to zero throughout the entire span, indicating that the average of family income over the beginning and end years of each period is only slightly more equal than the beginning-year measure.}

\footnote{16 Hungerford (2011) reports values of Fields mobility-as-equalizer measure (measured over all years, not just beginning and end—see appendix discussion), showing increases between the 1980s and 1990s; my results also show a rise between 1979–1989 and 1989–1999 (Figure 3). Hungerford also reports the Gini mobility measure as declining by a statistically significant amount (from 0.36 to 0.34) between the 1980s and 1990s, whereas my Gini mobility measure decreases by even more (from 0.38 to 0.32) between 1979–1989 and 1989–1999.}
mobility indicates the percentage of all individuals moving out of (up or down from) their quintile of origin defined in real dollar terms. This fraction rose somewhat from the 1970s through the 1980s, declined slightly, and then held steady.

**Origin-specific dollar-relative mobility**

Figure 4 also plots measures of real dollar movements by those beginning the period in the poorest or richest quintiles. Dollar movement out of the poorest quintile involves an increase in real income large enough to move above the constant-dollar beginning-of-period quintile ceiling by the end of the period (“percentage of the poorest who rise past real ceiling”). Reflecting overall real income growth as well as relative starting position, these dollar-relative mobility fractions are larger than the position-relative ones shown for the poorest quintile in Figure 2. Indeed, the time pattern of the percent poor climbing above the real dollar ceiling of their origin quintile tracks very closely the time pattern of within-period real income growth (the line immediately below it in Figure 4); the corresponding percentage of the richest quintile falling below the real dollar floor of their origin quintile displays roughly its mirror image. Thus, absolute mobility out of the poorest quintile rises steeply from 1972–1982 to 1982–1992 as real income growth increases, declines slightly, and then moves more or less sideways; the reverse is true of absolute mobility out of the richest quintile (except that the low point is 1981–1991, the peak period for real income growth). The dashed lines plot the percentages of individuals in the poorest/richest quintiles who move far in absolute terms—beyond the real dollar ceiling/floor of the adjacent quintile. Not surprisingly, real income growth augments dollar-relative upward mobility for the poor and suppresses dollar-relative downward mobility for the rich; since real income growth bears little relationship to the position-relative movements of these same poor and rich, the time patterns of the position-relative mobility measures differ from these dollar-relative measures.

Earlier research investigating these absolute quintile mobility measures is mixed. Acs

17 The solid black line plots the average change (across all families) in the log of family income between the beginning and endpoint of the period, that is, average real income growth during the period. This is the “directional income movement” measure as defined by Fields and Ok (1999), which I do not report as a mobility measure because it has no relative aspect.
and Zimmerman (2008b) report a drop in overall absolute quintile mobility from 0.627 to 0.598 between 1984-94 and 1994-04 (not significantly different from zero), a near-zero increase in mobility out of the bottom quintile and a statistically significant increase in mobility out of the top quintile (their Table 4). Hungerford (1993, combining his data for deciles into quintiles) shows overall absolute quintile mobility falling slightly (from 60.5 to 59.5) between the 1970s and 1980s, with absolute upward mobility from the poorest quintile falling, and absolute downward mobility from the richest quintile increasing.

**Overall dollar-relative mobility, part 2: Shorrocks M**

Shorrocks M compares the inequality of family income averaged over several years (the “long term”) with the inequality of single-year family income averaged over the years included in the long term. Shorrocks-M mobility is higher when individuals’ year-to-year movements (reflected in the long-term average) offset a greater fraction of the inequality of single-year incomes. Any relative mobility (families changing places in the distribution) makes the distribution of long-term income more equal than the (average) distribution of single-year incomes during the period; Shorrocks M quantifies the degree to which it is more equal. Long-term income is defined as the average of all the observations on each person’s family income within a 10-year period. The inequality of that long-term income is computed using three alternative measures: the Gini coefficient, Theil entropy, and the mean log deviation (MLD).

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18 Acs and Zimmerman (2008b) is the most comparable analysis of recent data in the literature; two key differences are the age range for their sample and their income definition. As I do, they use the PSID, employ two-year average endpoints, examine mobility in a 10-year period, and sample only heads and spouses (partners); the period that they call 1994–2004 uses an average of 1992 and 1994 income for the beginning of the period and averages income in 2002 and 2004 for the end (this is the period that I call 1993–2003). They base their analysis on a pre-tax, post-transfer definition of income, while I use post-government (post-tax, post-transfer) income; they limit their study to heads and partners who are between 25 and 45 years old at the beginning of the period, while my age limit of 16–62 becomes 16 to 52 in their beginning-of-period terms; hence I include people from nine years younger through seven years older than they do. They adjust for family size using the PSID-provided USDA needs measure, while I use the square root of family size.

19 As noted earlier in the text, available data limit the computation to income averaged over the six years comprising every other year within each 10-year period. However, for the periods from 1969–1979 through 1986–1996, data are available for every year, making possible a robustness check: Averaging 11 rather than six years of income yields a measure of inequality that is smoother than the every-other-year measure, but almost identical in level and time pattern.
Figure 5 displays Shorrocks M based on the three inequality measures (also Table 4). All three show similar time patterns, declining until the 1981–1991 period, increasing thereafter, and ending with higher levels of Shorrocks-M mobility than in the first few periods. The largest increase occurs for the MLD-based measure, which indicates that mobility offset about 27 percent of short-term inequality in the first few periods and 32 percent in the last few. By contrast, the Gini-based measure shows a lower level of mobility and very little increase.20

Although mobility has offset an increasing fraction of single-year inequality, a key background fact is that the U.S. income distribution has displayed steep increases in single-year or cross-sectional family income inequality in the 1980s, 1990s, and 2000s, as shown in Figure 6. An extensive literature has documented this rise in family income inequality; for example, Burkhauser, Feng, and Jenkins (2009) and Gottschalk and Danziger (2005) report steep increases in family income inequality in the 1980s, with further increases, but at a slower pace, in the 1990s.21 In the context of rising inequality, an alternative way to evaluate mobility over time would be to compare it with a benchmark in which mobility was sufficient to prevent long-term inequality from increasing (instead of Shorrocks’ implicit benchmark of offsetting a constant fraction of short-term inequality). Any mobility makes the long-term distribution of income more equal than the short-term distribution and most observers would agree that it is a good thing that mobility offset a greater fraction of short-term inequality in recent periods than in earlier periods; a logical extension of this argument during an era when short-term inequality has risen steeply is that it would be an even better thing if mobility were preventing long-term inequality from rising or were even causing long-term inequality to fall.

20 Transitory shocks are likely to show up as extreme values in a single year—hence contribute disproportionately to single-year inequality—and be wiped out by long-term averaging. The Theil and MLD measures are more sensitive to extreme income values (top and bottom, respectively) than is the Gini; as a result, they show more reduction in inequality than the Gini in moving from short-term to long-term income. Over time, the MLD-based mobility measure would rise more than the Gini if, for example, year-to-year (transitory) income changes that disappear in the long-term average, especially at the bottom of the distribution, accounted for an increasing fraction of (rising) single-year inequality. Hungerford (2011) reports an increase in the Gini-based Shorrocks-M mobility (long-term equalization) measure, from 0.109 in the 1980s to 0.111 in the 1990s; mine rises from 0.126 to 0.131 between 1978–1988 and 1988–1998.21 Burkhauser et al. (2009) actually find that inequality rose and then fell in the 1990s using the P90/P10 measure, but ended the decade higher than it began; using the Gini coefficient, they find steadier increases in the 1990s.
Thus, mobility might be said to increase when long-term inequality declines and be judged to fall when long-term inequality rises. According to this criterion, mobility has declined markedly in the United States over the last three decades, as long-term (10-year average) incomes have grown increasingly unequal—see Figure 7 and the right-hand panel of Table 4. According to all three measures, the inequality of 10-year average post-government income rose slightly between 1969–1979 and 1973–1983, increased markedly through 1985–1995, dipped briefly and then rose again, with an especially steep jump at the end of the data (between 1994–2004 and 1996–2006). The rise in all the measures of long-term inequality from the 1970s to the 1990s–2000s indicates directly that mobility across the 1969–2006 time span was not sufficient to offset the substantial rise in cross-sectional (one-year) inequality, leaving a greater gap between poor and rich families, even considering 10 years of post-government income. Indeed, even using 16-year periods, the inequality of long-term (16-year average) family incomes rose considerably according to all three inequality measures between the 1969–1985 period and the 1990–2006 period (data not shown).

**Time Trends**

The time patterns of position-relative mobility measures displayed in Figures 1 and 2 suggest general declines in mobility over time, especially since the 1980s. Many of the measures can be estimated directly with a regression based on the individual observations; pooling individual observations from two periods, clustering for overlap, and estimating the coefficient on a period-interaction term generates a test statistic for the significance of the difference between mobility rates in the two periods. Alternatively, bootstrap standard errors can be used to test the significance of differences between periods.

Many of the position-relative measures peaked in or near the 1977–1987 period. The difference between mobility during 1977–1987 and during the (most recent) 1995–2005 period is negative—mobility is lower in the more recent period—and significantly different from zero for all five of the position-relative overall measures shown in Figure 1 (p<0.01, except for rank

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22 Hungerford (2011) also reports a higher long-term Gini in the 1990s than in the 1980s.
Much of the earlier literature on mobility over time compares measures in two adjacent periods; thus, another way to evaluate the time patterns of our measures is to compare adjacent decades within our time span. For the position-relative measures shown in Figures 1 and 2, changes over eight adjacent (odd-year) decade pairs are available in the 1969–2006 time span. Among the four origin-specific position-relative measures shown in Figure 2, only the fraction of the poorest quintile moving “far” declined to a statistically significant degree between 1977–1987 and 1995–2005.

How do these decade-pair results compare with findings in the earlier decade-pair literature? Hungerford (2011) reports statistically significant declines in mobility between the 1980s (1979–1988) and 1990s (1989–1998). For example, 79 percent of individuals moved out of their decile of origin in the 1980s and 77 percent in the 1990s; the results here show a similar decline in decile mobility when comparing 1979–1989 and 1989–1999, but it is not statistically significant. Among the five overall measures, from six to eight of the eight pairs show declines in mobility between adjacent non-overlapping decades, but only two to four of the declines are statistically significant. Thus, for example, per capita decile movement has a lower value in the later of two adjacent decades in all eight comparisons, and the decline is significant in four of those cases, while M-hat is lower for the second decade in the pair in six out of eight pairs and only three of those declines are significantly different from zero. For the position-relative origin-specific measures, three to seven of the eight pairs show declines, but only zero to two of the declines between adjacent decades are statistically significant.

How do these decade-pair results compare with findings in the earlier decade-pair literature? Hungerford (2011) reports statistically significant declines in mobility between the 1980s (1979–1988) and 1990s (1989–1998). For example, 79 percent of individuals moved out of their decile of origin in the 1980s and 77 percent in the 1990s; the results here show a similar decline in decile mobility when comparing 1979–1989 and 1989–1999, but it is not statistically significant. While Hungerford (2008) finds 47 percent of the poorest quintile moving up in the 1980s and the 1990s, the fraction of the richest quintile moving down decreased from 50 percent to 46 percent and fewer of the rich (22 vs. 26 percent) moved far in the 1990s than in the 1980s; the data here similarly show a (statistically significant) decline in rich individuals moving far

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23 The difference between rank correlation’s slightly later peak period (1978–1988) and 1995–2005 is significantly different from zero with p<.01.

24 This analysis uses only odd-year pairs to avoid over-emphasizing the first half of the time span where the data include both even and odd-year period endpoints.

25 Hungerford uses post-government income adjusted for family size with a measure of poverty or needs, while I use two-year average endpoints and adjust using the square root of family size.
when comparing 1979–1989 with 1989–1999, but also show a significant decline in poor individuals moving far. Acs and Zimmerman (2008a) show modest declines in mobility between 1984–1994 and 1994–2004 (equivalent to this paper’s 1983–1993 and 1993–2003 periods), which are generally not statistically significant. For example, M-hat computed from their mobility matrices is 0.75 in 1984–1994 and 0.74 in 1994–2004.26 They report the percentage of heads and partners who move up from the bottom quintile as 44.4 in 1984–1994 and 39.0 in 1994–2004, and the fraction who move far from the bottom quintile (beyond the second quintile) as 20.3 percent and 15.7 percent, respectively, in the two periods, with none of these changes statistically significant. Table 2’s results for overall quintile mobility (M-hat) and far moves by the poor, like Acs and Zimmerman’s, fail to show statistically significant declines between 1983–1993 and 1993–2003, notwithstanding the fact that these measures (and others) show statistically significant declines when compared over the longer span, 1977–1987 to 1995–2005.27

Because real income growth peaked in the 1981–1991 period, many of the dollar-relative measures began falling only after that, so the interesting question is whether dollar-relative mobility measures declined significantly between 1981–1991 and 1995–2005. The degree to which end-year income reflects beginning-year income (income elasticity) was higher (mobility lower) in 1995–2005 than in 1981–1991, but not significantly so; the same was true for Fields mobility-as-equalizer. Gini mobility declined by a statistically significant amount between 1981–1991 and 1995–2005, as did the fraction of all individuals moving out of their absolute quintile groups during a period, although the significance levels are weaker than for the position-relative measures (Table 5).

Time trends for the origin-specific dollar-relative measures vary depending on the origin group: the fraction of the poorest quintile who moved above that quintile’s constant-dollar ceiling by the end of the period declined significantly between 1981–1991 and 1995–2005

26 Their Table 2, page 22.
27 As noted earlier in the text, most of the other studies that examine changes in mobility over adjacent long periods, including Gittleman and Joyce (1999), Hungerford (1993), Sawhill and Condon (1992), and Gottschalk and Danziger (2001), compare the 1970s and 1980s; most of the position-relative and dollar-relative mobility measures in this paper did not decline between 1970–1980 and 1980–1990, even though many did decrease significantly across more recent decades.
(p=0.06), as did the fraction of the poor who moved far (beyond the adjacent absolute quintile, 
p=0.03). By contrast, as Figure 4 suggests, the fraction of the richest quintile falling below the 
absolute floor of their origin quintile, or falling beyond the adjacent quintile’s constant-dollar 

A statistical test is unnecessary to see that the Shorrocks M mobility measures (Figure 5) 
have not declined; they are higher in 1995–2005 than earlier. However, the key mobility 
outcome—long-term inequality—shows sizable increases (Figure 7); indeed, even-year adjacent 
1996–2006) show statistically significant increases in inequality in all pairs except one (1982– 
1992–2002) for all three long-term inequality measures, and hence highly significant increases 
over the full time span. These increases in inequality indicate a substantial decline in mobility, 
in the sense that year-to-year income movements have been insufficient to offset fully the 
increasing within-year disparities, leaving U.S. families’ long-term income prospects diverging.

In summary, most of the position-relative mobility measures and a couple of the dollar-
relative measures show statistically significant declines in mobility between the 1980s and the 
most recent 1995–2005 period. As other researchers have noted repeatedly, even unchanged 
 mobility leads to widening inequality of long-term incomes as short-term inequality increases. 
The significant increase in long-term inequality documented here is further evidence reinforcing 
the conclusion that year-to-year changes in families’ incomes have become less effective in 
altering their long-term prospects—their position in the distribution in any one year is an 
increasingly good predictor of their position during the ensuing 10 years or at the end of that 
period.

**Pre-Government to Post-Government Income—the Effects of Taxes and Transfers**

The base case employs a post-government measure of income adjusted by the square 
root of family size. Changes over time in mobility based on adjusted post-government income 
reflect a combination of changes in the (government’s) tax and transfer system and changes in
the extent of individuals’ pre-government family income movements. The direct effects of government taxes and transfers on family incomes can be summarized by applying “mobility” measures to transitions from pre-government to post-government income in each year, indicating the extent of position-relative or dollar-relative movement from individuals’ pre-government family income positions to their post-government positions. For example, about one-quarter of heads and spouses are in a decile of the post-government adjusted family income distribution that differs from their decile in the pre-government distribution.

Such measures indicate an increasingly redistributive effect of the tax and transfer system on post-government incomes between 1969 and 1981, followed by (non-monotonic) declines through 2006; see Figure 8. For example, the fraction of individuals in a different decile or quintile of the post-government income distribution than in the pre-government distribution rose across the pre-1981 years and declined thereafter, as did the inverse rank correlation between the two distributions (increasing numbers/extent of changes in rank between the pre- and post-government income distributions in early years and declines later). Interestingly, more of these position-relative decile or quintile changes occur near the bottom of the distribution than near the top, with, for example, “far” moves (to the middle or lower quintiles) by the richest quintile equal to zero in most years and never more than 0.7 percent. In addition to redistribution, the tax and transfer system created the biggest gap between mean pre- and post-government incomes in 1981 (with mean adjusted post-government income 26 percent lower than mean adjusted pre-government income).

Similarly, the dollar-relative measures (Fields mobility-as-equalizer, Gini mobility, and inverse income elasticity) show increasing and then decreasing differences in income positions
between the two distributions.\textsuperscript{31} The inequality (as measured by the Gini coefficient) of the post-
government distribution is lower than that of the pre-government distribution (reflecting, of
course, one of the purposes of the tax and transfer system, and consistent with U.S. Census
Bureau (2007) findings of Gini coefficients of 0.49 for market income and 0.42 for disposable
income in 2005); the inequality of both rose quite markedly over the 1969–2006 time span, but
the \textit{difference} between the pre- and post-government distributions’ Gini coefficients increased
between 1969 and 1982 and fell thereafter.

One of the most important determinants of the post-government family income
distribution is the federal tax system, the effects of which are estimated (by Cornell University
researchers in their CNEF data set) using the NBER’s TAXSIM model. While TAXSIM is the best
method available, the researchers made two assumptions that reduce its accuracy in estimating
actual tax payments: (1) all taxpayers take the standard deduction; (2) everyone eligible for the
earned income tax credit (EITC) files for and receives it. The standard deduction is a more
questionable assumption at higher levels of income, while the EITC comes into play at the
bottom of the distribution. The EITC’s coverage has expanded considerably—both in terms of
dollars of benefits and in terms of number of eligible families—beginning in 1986 and
continuing sporadically to the present. Because the EITC is refundable, some eligible families
would not otherwise be filing income tax forms and hence might not file for the EITC; a
literature review estimated that 15 to 25 percent of eligible taxpayers fail to claim the credit
(Holt, 2006 and 2011). Because of its expansion, the EITC has increasingly augmented the post-
government incomes of low-income families, especially those with children, and has thereby
contributed to upward mobility during the periods spanning the expansions; however, the
assumption that all eligibles receive the credit implies that such upward mobility is somewhat
overstated in our mobility estimates.\textsuperscript{32}

\textsuperscript{31} Again not surprisingly, the absolute quintile measures show the tax and transfer system moving a
much larger fraction (63 percent) of the richest pre-government quintile below the quintile’s (pre-
government) real dollar floor than the fraction (7 percent) of the poorest pre-government income quintile
raised by the tax and transfer system above that quintile’s dollar ceiling.

\textsuperscript{32} Indeed, the 1986 bump in Figure 8 may reflect a jump in TAXSIM’s estimated EITC payments with the
1986 expansion. By the same token, the assumption that all filers use the standard deduction would
The declines in redistribution and overall taxes after 1981 presumably reflect the Reagan tax cuts enacted in the Economic Recovery Act of 1981 and continuing intermittently through the Bush tax cuts of 2001–2003. Indeed, the pattern of increasing and then declining influence of the tax and transfer system is consistent with estimates by Piketty and Saez (2006), which show rising average federal tax rates (individual income and payroll taxes) from 1960 through 1981, followed by modest declines in rates.

Comparing Mobility in Pre-Government Income with Base-Case Post-Government Income

Measures of 10-year mobility based on pre-government income differ to varying degrees in both levels and time trends from those based on post-government income. For position-relative measures, mobility based on pre-government income is generally lower than that based on post-government income—especially during the late 1970s–1980s (for example, 1976–1986 through 1980–1990), when many of the post-government measures peaked—and the pre- and post-government measures have moved closer together in the most recent periods (Figure 9). Overall, the position-relative measures for pre-government income are somewhat more level (rather than increasing) before the 1980s, and they show smaller declines from the 1980s to the most recent period than those shown in Figures 1 and 2 for post-government income.

The story is different for overall dollar-relative mobility measures. While Gini mobility is similar for pre- and post-government income, both the inverse income elasticity and Fields mobility-as-equalizer are noticeably lower in the early periods for pre-government than for post-government income and then either catch up to the corresponding measure for post-government income (the Fields measure) or climb steeply beyond it (inverse income elasticity). As a result, neither displays the declines in mobility for pre-government income between 1970–

increasingly understate the post-government incomes of higher-income taxpayers if itemization were increasing in or becoming more concentrated in the upper quintiles of the distribution, but IRS Statistics of Income data do not show steady increases or decreases in itemization during the 1969–2006 time span.

33 The comparisons of pre-government mobility with post-government mobility in this text discussion and Figures 9 and 10 are based on an “overlap sample” of observations—those that are present in both the pre-government and post-government income samples—that were separately trimmed to eliminate outliers, as described earlier.
1980 and 1980–1990 that were visible for post-government income in Figure 3, although they do decrease between 1993–2003 and 1995–2005 as post-government mobility does. Thus, when taking account of dollar as well as relative moves, pre-government income mobility rose markedly between the 1970s and the 1980s, producing a picture of generally increasing mobility over the full time span, or at least through 1993–2003.

Overall absolute quintile mobility is slightly lower for pre-government income than for post-government income after the 1977–1987 period. Data on moves by those who begin in the richest and poorest quintiles show lower absolute pre-government income mobility for the poor and higher absolute pre-government income mobility for the rich (after 1979–1989), and the same is true for “far” absolute moves by these two quintiles. These differences presumably reflect the positive effects of the tax and transfer system in making the poor more upwardly mobile across the real income ceiling of their starting quintile, and protecting the rich against downward moves across the real income floor of their starting quintile. Because these level differences between pre- and post-government absolute mobility have increased over time, absolute mobility shows a decline over the full time span (and since 1981–1991) for the poorest pre-government quintile (both any moves up and far moves) and a rise for the richest (Figure 10).

For Shorrocks M, pre-government mobility is slightly lower than post-government mobility based on the Gini and Theil—that is, year-to-year income changes within a 10-year period offset a slightly lower fraction of short-term inequality of pre-government income than of post-government income; the opposite is true—measured mobility is much higher for pre-government incomes—when inequality is measured using the mean log deviation. \(^{34}\) Over time, Shorrocks M measures based on all three inequality indexes declined somewhat more between 1970–1980 and 1981–1991 and then rose less from the 1980s through the most recent period for pre-government than for post-government income.

Looking directly at long-term inequality as an indicator of mobility, all three measures

\(^{34}\) Recall that the pre-government distribution includes non-negligible numbers of families with zero incomes, whose post-government incomes consist only of government transfers. The sensitivity of the MLD to inequality at the bottom of the distribution produces extremely high measured inequality of the pre-government distribution, both in a single year and long-term.

One possible explanation for these pre- and post-government time patterns is that the rising redistributive impact of taxes and transfers during the first dozen years of the 1969–2006 time span (through 1981—recall Figure 8) combined with increases in (post-government) real income growth through the 1981–1991 period contributed to rising post-government income mobility according to a variety of measures across the periods that straddled the rise in redistribution: Most of the position-relative measures peaked around 1977–1987; some of the dollar-relative measures (notably, absolute quintile mobility and the fraction of those starting in the poorest quintile who moved up across the real dollar quintile ceiling or moved far) were highest around the 1981–1991 period as the pace of growth in real post-government income also peaked. At the same time, government redistribution presumably slowed the rise in long-term inequality of post-relative to pre-government income through 1981–1991. As the redistributive role of government lessened after 1981 and real income growth slowed and steadied, family income mobility declined sporadically, long-term inequality increased more for post-government than for pre-government incomes (through 1992–2002), and almost all the mobility measures declined across the most recent periods.

Summary and Discussion

A variety of measures indicate that U.S. family income mobility has decreased over the 1969–2006 time span, and especially since the 1980s. Most overall relative mobility measures fell, on net, with mobility significantly lower in the 1995–2005 period than in the 1977–1987 period. Origin-specific measures also declined, with families starting in the poorest quintile less likely to make far moves (to the middle or higher quintiles) during the 10 years from 1995 to
2005 than within the 1977–1987 period. Similarly, absolute quintile mobility declined overall and for poor families after 1981–1991 (when real post-government income growth began to slow). Thus, families have become increasingly less likely to change rank or move out of their (relative or absolute) decile or quintile of origin, and when they move, to move less far; this appears to be the case overall and for those who start rich or poor. The direct measure of mobility’s equalizing effects (Shorrocks M, the fraction of single-year inequality that is offset by mobility), shows higher mobility in the 1990s and into the 2000s, when single-year inequality was higher, than in the 1970s and ‘80s.

Between the earliest and the latest periods, many of the measures move up over several periods, and then down, or vice versa—the downward trend is not at all monotonic. This is undoubtedly part of the reason that most earlier research failed to document a downward trend; that research nonetheless noted that even steady levels of mobility imply increasing long-term inequality during the last quarter of the 20th century and into the 21st when cross-sectional inequality in the United States was rising—increases in long-term inequality that are documented here. That is, when judging mobility by its outcome—the inequality of long-term income—the verdict is decreasing mobility, as poor and rich families have grown increasingly far apart, even considering average income over 10 (or even 16) years. Furthermore, all the mobility measures— even Shorrocks M—show a decline in mobility between the next-to-last 10-year period and the most recent: mobility was lower between 1995 and 2005 (or 1996–2006) than between 1993 and 2003 (or 1994–2004). These recent data have become available since previous mobility research was published, providing additional cause for concern.

Increasing economic mobility is a widely shared goal, especially if mobility equalizes opportunity and lifetime incomes or takes the form of families at the bottom moving up (Economic Mobility Project, 2009); against this goal, this paper’s findings are discouraging. Long-term income adjusted for family size is considerably more unequally distributed among families for periods ending in the 2000s than for periods ending in the 1970s and ‘80s. That is, family income mobility has been insufficient to stem increases in inequality of long-term income. Furthermore, other mobility measures indicate that a family’s position at end of a period in the 2000s was less likely to have been produced by a random process and more
correlated with its start position than was the case 20 years earlier.

What are the implications for policy? It appears that increasingly redistributive U.S. tax and transfer policies may have enhanced family income mobility from the 1970s into the 1980s, but then had decreasing impact as tax rates were reduced after 1981 (and intermittently thereafter). Beyond overall patterns, the data indicate that the typical individual in the poorest one-fifth of the family income distribution is less likely to move up beyond that group’s real dollar ceiling within a decade than it was 15 to 20 years earlier. These facts suggest that policy remedies for those at the bottom should aim beyond short-term help, as those who are poor at any point in time are now likely to have low long-term incomes. Beyond this, the choice of policy presumably hinges, at least in part, on the reasons for the decline in mobility, for example, whether it reflects rising barriers to opportunity, other shifts in the economy, or changes in the U.S. tax and transfer system that increasingly reinforce rather than offset market disparities over time.

Further research is needed to assess the balance among these potential sources of the decline in mobility. The first step will be to investigate time-series determinants of mobility patterns during the 1969–2006 period, examining the degree to which shifting business cycle conditions are associated with higher or lower mobility.
Appendix: Specific Measures of Mobility

The mobility measures reported and analyzed in this paper are categorized in Table 1. This appendix discussion proceeds from the position-relative overall measures in the upper left quadrant of Table 1 to position-relative origin-specific measures (upper right quadrant), through the dollar-relative overall and origin-specific measures (lower left and lower right quadrants).

Most of the overall, position-relative mobility measures examined in the paper are based on a transition matrix $P$ for the family income distribution divided into quintiles or deciles, in which the matrix cells $p_{ij}$ (indexed by row $i$ and column $j$) represent the fraction of all individuals in the entire distribution who start a period in quintile or decile $i$ and end the period in quintile or decile $j$. Alternatively, they are based on the row mobility matrix $Q$, in which the matrix cells $q_{ij}$ represent the fraction of those individuals who start a period in quintile or decile $i$ who end the period in quintile or decile $j$. To clarify the distinction, all the $i \times j$ cells in the transition matrix $P$ sum to one, while the $j$ elements in each row of the mobility matrix $Q$ sum to one.

Decile mobility is the fraction of individuals who move up or down at least one decile. Thus it is equal to one minus the fraction of individuals along the diagonal of the decile transition matrix:

$$1 - \sum_{k=1}^{10} p_{kk} \quad \text{or} \quad 1 - \text{trace } P .$$

(1)

Per capita decile movement, the average distance (in deciles) that individuals move during a period, contains somewhat more information about off-diagonal elements of the transition matrix:

$$\frac{1}{N} \sum_{i=1}^{N} | \text{decile}(y_{ie}) - \text{decile}(y_{ib}) | ,$$

(2)

where $y_{ib}$ and $y_{ie}$ are family income (or another measure of wellbeing, such as family income adjusted for family size, the measure used in this paper) of the $i^{th}$ individual in the beginning ($b$) and end ($e$) years of a period, and $N$ is the number of individuals whose incomes are observed at both the beginning and end of a period.
Shorrocks (1978b) developed a measure scaled to vary between zero and one that is based on the trace of the mobility matrix \( Q \):

\[
M\text{-}\hat{\text{m}} = \frac{k - \text{trace } Q}{k - 1},
\]

where \( Q \) is the \( k \times k \) row mobility matrix reporting the fraction of those who start in each beginning-of-period quintile or decile who end in each end-of-period quintile or decile. In this paper, \( M\text{-}\hat{\text{m}} \) is calculated for the quintile matrix \( (k = 5) \) to facilitate comparison with the results of other researchers who publish their quintile mobility matrices.\(^1\) \( M\text{-}\hat{\text{m}} \) reaches its maximum value of one when those who start in each quintile have an equal probability of ending in all quintiles, implying that all rows of the matrix are identical; this is Shorrocks’ definition of “perfect mobility.” \( M\text{-}\hat{\text{m}} \) equals zero (no mobility) when everyone’s end-of-period quintile is the same as the one in which he or she began; the only non-zero entries in the matrix are ones on the diagonal.

One of the overall position-relative measures is based on rank, not deciles or quintiles. The rank correlation (inverse) — the simplest indicator of time independence — shows how strongly beginning-of-period position is related to end-of-period position. It is computed as one minus the correlation between individuals’ income ranks at the beginning and end of the period.\(^2\)

A final overall position-relative measure is the adjusted contingency coefficient \( C^* \), based on the chi-squared test-statistic related to independence of rows and columns in the decile transition matrix\(^3\) and “adjusted” to range between zero and one. Subtracting \( C^* \) from one makes measured mobility higher when the association between starting and ending decile is lower:

\(^1\)Except that it is based on quintiles instead of deciles, the information content of \( M\text{-}\hat{\text{m}} \) is the same as that of decile mobility — both depend on the diagonal elements of the mobility or transition matrix, that is, the traces of those matrices.

\(^2\)Subtracting from one rescales the measure to be higher when mobility is higher. Note that a negative correlation would cause the rescaled mobility measure to exceed one. Conceptually, treating one as the maximum value for the inverse rank correlation is equivalent to Shorrocks’ rescaling in the calculation of \( M\text{-}\hat{\text{m}} \), in that it defines “perfect (or maximum) mobility” as the situation in which end-of-period rank bears no relationship to beginning-of-period rank. As Burkhauser and Couch (2009) note, the “conflict in how to norm... the upper value of a mobility index is a recurring theme in this [mobility] literature” (p. 529). A negative correlation would indicate even greater changes in rank than a zero correlation, but does not represent what Shorrocks calls a realistic case. And in fact negative correlations never occur in the data analyzed here.

\(^3\)The contingency coefficient \( C^* \) can be computed only for matrices with an equal number of rows and columns.
\[ 1 - \text{decile C}^* = 1 - \frac{\sqrt{\chi^2}}{\sqrt{\frac{N + \chi^2}{k}}} \tag{4} \]

where \( N \) is the total (weighted) number of individuals, \( k \) is 10 — the number of rows (columns), and \( \chi^2 \) is the chi-squared test statistic:

\[ \chi^2 = \sum_{i,j=1}^{k} \frac{(n_{ij} - e_{ij})^2}{e_{ij}} \tag{5} \]

with \( n_{ij} \) the observed number (not percentage) of people who start in decile \( i \) and move to decile \( j \) during the period and \( e_{ij} \) the number that would be expected based on the product of row and column totals \( N_i \) and \( N_j \), respectively, (that is, if beginning- and end-of-period decile position were completely independent).

Several position-relative origin-specific measures (upper right quadrant of Table 1) are based on the quintile row mobility matrix \( Q \). These include the fraction of individuals starting in the poorest (richest) quintile who move up (down), equal to one minus the upper left (lower right) corner elements in the quintile row mobility matrix:

\[ 1 - q_{11} \quad \text{and} \quad 1 - q_{55} \tag{6} \]

and the fraction of individuals beginning in the poorest (richest) quintile who move beyond the adjacent quintile:

\[ \sum_{j=3}^{5} q_{1j} \quad \text{and} \quad \sum_{j=1}^{3} q_{5j} . \tag{7} \]

These latter measures are used fairly widely in the literature to indicate the fraction of those beginning in the poorest and richest quintiles who move “far,” that is who move up (down) more than one quintile.

As noted in the text, dollar-relative mobility measures (lower panel of Table 1) involve changes in both income and relative position in the distribution of income.

Gini mobility (Yitzhaki and Wodon, 2005) reflects changes in relative rank and real income. It compares the covariance of individuals’ changes in income and rank over the period with the
covariances of income and rank at the start and the end of the period:

\[
\sum_{i=1}^{N} \left( \frac{y_{ie}}{\mu_e} - \frac{y_{ib}}{\mu_b} \right) \left( \frac{R_{ie}}{N} - \frac{R_{ib}}{N} \right) \]

where \( \mu_b \) and \( \mu_e \) are mean incomes at the beginning and end of the period, and \( R_{ib} \) and \( R_{ie} \) are the rank of individual \( i \) in the beginning- and end-of-period distributions, from poorest (lowest rank) to richest (highest rank). Yitzhaki and Wodon developed their measure to provide (along with the Gini index of inequality) “an overall consistent framework for the analysis of mobility [and] inequality”; they cite examples of job rotations and the distribution of income over the life cycle, in which snapshot measures of inequality may give misleading indications of longer-term inequality that their Gini mobility measure will correct. Note from the numerator expression that if either an individual’s relative income or his/her rank does not change, the contribution to Gini mobility is zero.

The elasticity of end-of-period income with respect to beginning-of-period income is the most widely used measure of intergenerational income mobility (typically used in examining the dependence of a man’s earnings on his father’s earnings at a similar life-stage). The elasticity is the estimated coefficient in a regression of the logarithm of end-of-period income on the logarithm of beginning of-period income. Subtracting the elasticity from one makes the measure higher when mobility is higher — end-of-period income is less tightly tied to beginning-of-period income.

The Fields (2008) mobility-as-equalizer statistic measures whether the average (or sum) of beginning-of-period and end-of-period income is more or less equal than beginning-of-period income:

\[
\text{Fields MAE} = 1 - \frac{I(y_{ib} + y_{ie})}{I(y_{ib})},
\]

where \( y_{ib} \) and \( y_{ie} \) are beginning- and end-of-period income for individual \( i \) and \( I(.) \) is the Gini (or other) index of inequality across all incomes. This measure has a directional aspect, in that

---

4They call it the Gini index of mobility because the Gini index of inequality is proportional to the covariance of income and rank.
it indicates the degree to which mobility equalizes average income compared with beginning-year income. The index takes on larger positive values (closer to one) when average incomes are more equally distributed than base-year incomes and becomes negative when the inequality of end-year income is so much greater than that of beginning-year income that average of those two incomes is less equally distributed than beginning-year income (that is, end-year inequality overwhelms the equalizing power of averaging two years of income).\footnote{A later paper (Fields et al., 2007) includes all years in the period, not just beginning and end years in this measure, but it is difficult to interpret the directional aspect of the measure (which Fields stresses) when multiple post-base-year incomes are included; in addition, only the denominators would distinguish Shorrocks $M$ (see below) and Fields $MAE$ with the multi-year approach. Thus, the beginning- and end-year formulation of the Fields measure is used here.}

Absolute quintile mobility can be characterized via absolute transition and mobility matrices like their relative counterparts $P$ and $Q$ above, where $B_p$ is a 5 x 5 absolute quintile transition matrix consisting of elements $bp_{ij}$ indicating the fraction of all individuals who begin in quintile $i$ and end in group $j$, where group $j$ is defined by the real income range of beginning-of-period quintile $j$. Similarly, absolute mobility matrix $B_q$ consists of elements $bq_{ij}$ indicating the fraction of individuals starting in quintile $i$ who end in group $j$. The absolute quintile mobility measure tallies the percentage of all individuals who move out of their beginning-of-period absolute quintile:

$$1 - \frac{1}{5} \sum_{k=1}^{5} bp_{kk},$$

(10)

Shorrocks (1978a) proposes a measure of “the degree to which equalization occurs as the observation period is extended”; it compares the inequality of long-term income (over $t$ years) with the (weighted average) inequality of single-year income (to produce a measure of rigidity, $R$) and scales it to increase with mobility by subtracting the ratio from one:

$$Shorocks\ M = 1 - R = 1 - \frac{I(y_{i1} + y_{i2} + ... + y_{it})}{\sum_{j=1}^{I} \mu_j \ast I(y_{ij}) / \sum_{j=1}^{I} \mu_j},$$

(11)

where $I(.)$, as above, is an index of inequality; $y_{i1}, y_{i2}, ..., y_{it}$ is the income of individual $i$ in year 1 through year $t$ (year 1 can be seen as year $b$, and year $t$ as year $e$). For this measure, three
alternative inequality measures are analyzed: the Gini, the Theil entropy measure, and the mean log deviation (MLD). Note that these are the only measures examined in this paper that use income data on the intervening years between beginning- and end-of-period. For this reason, the Shorrocks $M$ and its “ingredient” measure of long-term inequality (the numerator in the expression above) constitute the most direct measures of the degree to which mobility equalizes long-term incomes.

The origin-specific, dollar-relative measures tally moves out of absolute quintiles, based on the absolute quintile mobility matrix $B_q$, defined above. As is the case for the position- relative origin-specific mobility measures, the figures of interest are the percentage of individuals starting in the poorest quintile whose incomes rise enough to move up past the real dollar upper bound of the poorest quintile and, similarly, the percentage of those beginning in the richest quintile whose incomes fall enough to end below the real dollar lower bound of the richest quintile:

$$1 - b_{q11} \quad \text{and} \quad 1 - b_{q55}.$$  \hfill (12)

Similar to the relative quintile measures, “far” moves by those beginning in the poorest and richest quintiles — beyond the constant-dollar boundaries on the far side of the adjacent quintiles — are also examined:

$$\sum_{j=3}^{5} b_{q1j} \quad \text{and} \quad \sum_{j=1}^{3} b_{q5j}.$$  \hfill (13)
References


Table 1. Measures of mobility analyzed in this paper

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<tr>
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<th>OVERALL</th>
<th>ORIGIN-SPECIFIC</th>
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<td>• Decile mobility</td>
<td>• Fraction of poorest (richest) quintile who move up (down)</td>
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<td></td>
<td>• Per capita decile movement</td>
<td>• Fraction of poorest (richest) quintile who move “far” -- to middle or farther</td>
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<td></td>
<td>• Shorrocks M-hat for quintiles</td>
<td>quintiles</td>
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<tr>
<td></td>
<td>• Rank correlation (inverse)</td>
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<td></td>
<td>• Decile adjusted contingency coefficient C* (inverse)</td>
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<td><strong>DOLLAR-RELATIVE</strong></td>
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<td></td>
<td>(inverse)</td>
<td>• Fraction of poorest (richest) quintile who move “far” in real terms --</td>
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<td>• Absolute quintile mobility</td>
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<td>entropy, and mean log deviation)</td>
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<td>• Inequality of long-term income</td>
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Note: See text for discussion and appendix for definitions and sources.
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<th>Shorrock's M-hat for quintiles</th>
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<th>Decile C* (inverse)</th>
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<th>% richest quintile move down</th>
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Notes: Base case uses post-government income adjusted with the square root of family size and two-year average endpoints. See appendix for definitions of measures. These data are displayed in Figures 1 and 2.
Sources: Author's calculations based on data from Panel Study of Income Dynamics and Cornell Cross-National Equivalent Files.
Table 3. Dollar-Relative Mobility Measures, base case

<table>
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<tr>
<th>Period:</th>
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<th>Income elasticity (inverse)</th>
<th>Fields MAE</th>
<th>Absolute quintile mobility (%)</th>
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Notes: Base case uses post-government income adjusted with the square root of family size and two-year average endpoints. See appendix for definitions of measures. These data are displayed in Figures 3 and 4.
Sources: Author’s calculations based on data from Panel Study of Income Dynamics and Cornell Cross-National Equivalent Files.
Table 4. Mobility Measured as Equalization of Long-Term Income, base case

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<tr>
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<td>13.6</td>
<td>25.9</td>
<td>27.7</td>
<td>0.220</td>
<td>0.078</td>
<td>0.082</td>
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<tr>
<td>1971-81</td>
<td>13.4</td>
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<td>27.7</td>
<td>0.219</td>
<td>0.076</td>
<td>0.081</td>
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<td>1972-82</td>
<td>13.8</td>
<td>26.5</td>
<td>28.4</td>
<td>0.221</td>
<td>0.078</td>
<td>0.083</td>
</tr>
<tr>
<td>1973-83</td>
<td>12.9</td>
<td>25.0</td>
<td>27.3</td>
<td>0.222</td>
<td>0.079</td>
<td>0.083</td>
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<tr>
<td>1974-84</td>
<td>12.9</td>
<td>24.8</td>
<td>26.7</td>
<td>0.228</td>
<td>0.083</td>
<td>0.088</td>
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<td>1975-85</td>
<td>12.7</td>
<td>24.8</td>
<td>27.3</td>
<td>0.230</td>
<td>0.084</td>
<td>0.089</td>
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<td>1976-86</td>
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<td>25.2</td>
<td>27.2</td>
<td>0.230</td>
<td>0.085</td>
<td>0.090</td>
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<td>25.1</td>
<td>27.2</td>
<td>0.232</td>
<td>0.086</td>
<td>0.092</td>
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<tr>
<td>1978-88</td>
<td>12.6</td>
<td>24.6</td>
<td>26.3</td>
<td>0.236</td>
<td>0.089</td>
<td>0.095</td>
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<tr>
<td>1979-89</td>
<td>12.5</td>
<td>24.6</td>
<td>26.7</td>
<td>0.240</td>
<td>0.092</td>
<td>0.098</td>
</tr>
<tr>
<td>1980-90</td>
<td>12.3</td>
<td>23.9</td>
<td>25.9</td>
<td>0.242</td>
<td>0.095</td>
<td>0.100</td>
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<td>1981-91</td>
<td>12.0</td>
<td>23.4</td>
<td>25.8</td>
<td>0.246</td>
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<td>0.104</td>
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<tr>
<td>1982-92</td>
<td>12.0</td>
<td>23.4</td>
<td>25.8</td>
<td>0.250</td>
<td>0.101</td>
<td>0.107</td>
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<tr>
<td>1983-93</td>
<td>12.4</td>
<td>24.2</td>
<td>26.7</td>
<td>0.252</td>
<td>0.103</td>
<td>0.109</td>
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<tr>
<td>1984-94</td>
<td>12.2</td>
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<td>26.5</td>
<td>0.254</td>
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<td>1985-95</td>
<td>12.1</td>
<td>24.0</td>
<td>26.4</td>
<td>0.256</td>
<td>0.105</td>
<td>0.113</td>
</tr>
<tr>
<td>1986-96</td>
<td>12.9</td>
<td>25.4</td>
<td>27.8</td>
<td>0.252</td>
<td>0.102</td>
<td>0.110</td>
</tr>
<tr>
<td>1988-98</td>
<td>13.1</td>
<td>26.1</td>
<td>28.7</td>
<td>0.255</td>
<td>0.104</td>
<td>0.112</td>
</tr>
<tr>
<td>1990-00</td>
<td>13.3</td>
<td>26.9</td>
<td>29.5</td>
<td>0.258</td>
<td>0.107</td>
<td>0.115</td>
</tr>
<tr>
<td>1992-02</td>
<td>14.3</td>
<td>28.4</td>
<td>31.8</td>
<td>0.258</td>
<td>0.107</td>
<td>0.117</td>
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<tr>
<td>1994-04</td>
<td>14.2</td>
<td>28.4</td>
<td>31.8</td>
<td>0.263</td>
<td>0.112</td>
<td>0.120</td>
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<tr>
<td>1996-06</td>
<td>13.5</td>
<td>27.7</td>
<td>32.0</td>
<td>0.278</td>
<td>0.125</td>
<td>0.135</td>
</tr>
</tbody>
</table>

Notes: Base case uses post-government income adjusted with the square root of family size. See appendix for definitions of measures. "Long-term" defined as 10-year period, averaging income over every other year within period. These data are displayed in Figures 5 and 7.

Sources: Author’s calculations based on data from Panel Study of Income Dynamics and Cornell Cross-National Equivalent Files.
Table 5. Significance tests for changes over time

<table>
<thead>
<tr>
<th>Measure</th>
<th>Type of estimate</th>
<th>Periods &amp; levels</th>
<th>Difference</th>
<th>S.E.</th>
</tr>
</thead>
<tbody>
<tr>
<td>Position-Relative Overall Mobility</td>
<td></td>
<td>1977-87  1995-05</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Decile mobility (%)</td>
<td>Logit</td>
<td>79.5  75.3</td>
<td>-0.240***</td>
<td>(0.077)</td>
</tr>
<tr>
<td>Per capita decile movement</td>
<td>OLS</td>
<td>1.92  1.72</td>
<td>-0.201***</td>
<td>(0.053)</td>
</tr>
<tr>
<td>Shorrocks M-hat for quintiles</td>
<td>Bootstrap</td>
<td>0.782  0.714</td>
<td>-0.068***</td>
<td>(0.021)</td>
</tr>
<tr>
<td>Rank correlation (inverse)</td>
<td>OLS</td>
<td>0.390  0.340</td>
<td>-0.049**</td>
<td>(0.024)</td>
</tr>
<tr>
<td>Decile C* (inverse)</td>
<td>Bootstrap</td>
<td>0.371  0.329</td>
<td>-0.043***</td>
<td>(0.016)</td>
</tr>
<tr>
<td>Position-Relative Origin-Specific Mobility</td>
<td></td>
<td>1977-87  1995-05</td>
<td></td>
<td></td>
</tr>
<tr>
<td>% poorest quintile move up</td>
<td>Logit</td>
<td>45.2  41.9</td>
<td>-0.133</td>
<td>(0.138)</td>
</tr>
<tr>
<td>% richest quintile move down</td>
<td>Logit</td>
<td>50.5  45.0</td>
<td>-0.222</td>
<td>(0.153)</td>
</tr>
<tr>
<td>% poorest quintile move far</td>
<td>Logit</td>
<td>21.2  15.4</td>
<td>-0.391**</td>
<td>(0.182)</td>
</tr>
<tr>
<td>% richest quintile move far</td>
<td>Logit</td>
<td>25.1  20.2</td>
<td>-0.281</td>
<td>(0.182)</td>
</tr>
<tr>
<td>Dollar-Relative Overall Mobility</td>
<td></td>
<td>1981-91  1995-05</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Income elasticity (inverse)</td>
<td>OLS</td>
<td>0.369  0.334</td>
<td>-0.035</td>
<td>(0.031)</td>
</tr>
<tr>
<td>Gini mobility</td>
<td>Bootstrap</td>
<td>0.286  0.246</td>
<td>-0.041*</td>
<td>(0.021)</td>
</tr>
<tr>
<td>Fields MAE</td>
<td>Bootstrap</td>
<td>0.0335  0.0189</td>
<td>-0.015</td>
<td>(0.016)</td>
</tr>
<tr>
<td>Absolute quintile mobility (%)</td>
<td>Logit</td>
<td>66.1  63.3</td>
<td>-0.123*</td>
<td>(0.065)</td>
</tr>
<tr>
<td>Dollar-Relative Origin-Specific Mobility</td>
<td></td>
<td>1981-91  1995-05</td>
<td></td>
<td></td>
</tr>
<tr>
<td>% poorest quintile rise past real ceiling</td>
<td>Logit</td>
<td>61.3  55.1</td>
<td>-0.252*</td>
<td>(0.135)</td>
</tr>
<tr>
<td>% richest quintile fall past real floor</td>
<td>Logit</td>
<td>20.8  23.0</td>
<td>0.127</td>
<td>(0.172)</td>
</tr>
<tr>
<td>% poorest quintile rise far in real terms</td>
<td>Logit</td>
<td>35.0  27.9</td>
<td>-0.333**</td>
<td>(0.150)</td>
</tr>
<tr>
<td>% richest quintile fall far in real terms</td>
<td>Logit</td>
<td>9.83  9.00</td>
<td>-0.098</td>
<td>(0.238)</td>
</tr>
</tbody>
</table>

Notes: Logit differences are expressed in terms of log odds.
Clustered standard errors are reported in parentheses (see text for details).
Asterisks denote significance as follows: * p<0.10, ** p<0.05, *** p<0.01
Sources: Author’s calculations based on data from Panel Study of Income Dynamics and Cornell Cross-National Equivalent Files.
Figure 1. Position-relative overall mobility measures
Base case

Note: Base case uses post-government income adjusted with the square root of family size and two-year average endpoints. See appendix for definitions of measures. These data are listed in Table 2.
Source: Author’s calculations based on Panel Study of Income Dynamics and Cornell University Cross National Equivalent File.
Figure 2.
Position-relative origin-specific mobility for poorest and richest quintiles
Base case

% of richest quintile who move down
% of poorest quintile who move far
% of richest quintile who move far
% of poorest quintile who move up

Note and source: see Figure 1.
Figure 3. Dollar-relative overall mobility measures

Base case

Note and source: see Figure 1. These data are listed in Table 3.
Figure 4. Absolute quintile mobility and income growth

Base case

Absolute quintile mobility %

% of poorest quintile who rise past real ceiling

Real income growth, right axis →

% of poorest quintile who rise far in real terms

% of richest quintile who fall past real floor

% of richest quintile who fall far in real terms

Note and source: see Figure 1. These data are listed in Table 3.
Figure 5.
Shorrocks M: Degree to which 10-year mobility offsets 1-year inequality
Base case

Note: Base case uses post-government income adjusted with the square root of family size and every-other-year income within 10-year period. See appendix for definitions of measures. These data are listed in Table 4.
Source: Author's calculations based on Panel Study of Income Dynamics and Cornell University Cross National Equivalent File.
Figure 6. U.S. household income inequality, CPS data

Note: Discontinuities between 1992 and 1993 reflect changes in data collection methods and increases in income limits for recording data.
Figure 7. Inequality of 10-year average income

Base case

Gini, right axis

Mean log deviation

Lower two lines, left axis

Note and source: see Figure 5.
Figure 8. Position-relative measures of overall "mobility" from pre-government income to post-government income

Per family decile movement

Decile mobility

M-hat

Decile C* (inverse)

Rank correlation (inverse)

Note: Mobility as “movement” from pre-government income position to post-government income position; incomes are adjusted using square root of family size. See appendix for definitions of measures.

Source: Author’s calculations based on Panel Study of Income Dynamics and Cornell University Cross National Equivalent File.
Figure 9. Position-relative mobility measures, pre-government income (overlap sample)

Note: Based on pre-government income adjusted with the square root of family size and two-year average endpoints. Overlap sample includes observations present in both pre- and post-government trimmed samples; see text. See appendix for definitions of measures.
Source: Author’s calculations based on Panel Study of Income Dynamics and Cornell University Cross National Equivalent File.
Figure 10. Absolute quintile mobility measures, pre-government income
(overlap sample)

**Absolute quintile mobility**

- % of poorest quintile who rise past real ceiling
- % of richest quintile who fall past real floor
- % of richest quintile who fall far in real terms
- % of poorest quintile who rise far in real terms

Note and source: see Figure 9.