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ABSTRACT

On Input Market Frictions and Estimation of Factors Demand^{*}

In this paper we explore the impact of imperfectly competitive input markets on production function estimation. First order profit maximizing conditions are altered when frictions in input markets cause the elasticity of input supply to the firm to be finite. A consequence of this is that the standard econometric model used for production function estimation will be misspecified. We prove that, in all non trivial cases, finite elasticities of supply to the firm will lead to inconsistent estimates of production function parameters. Monte Carlo simulations show that the resulting bias can be economically significant.

JEL Classification: D2, J42, J23

Keywords: input market frictions, labor market frictions, estimation of factor demand

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1 Introduction

The last fifteen years have seen a dramatic increase in the literature on labor market imperfections. Work by Burdett and Mortensen (1998) has established the role of search costs in creating labor market frictions. Manning (2003) shows how, even in the presence of many competing firms in the labor market, these frictions imply that a monopsony model may best represent the functioning of the labor market. Recent empirical tests have supported these works by uniformly rejecting the hypothesis of perfectly elastic labor supply to the firm¹.

Separately, many great contributions to our understanding of labor markets have come through empirical estimates of production function parameters. Studies on skill biased technological change and changes in relative wages include Katz and Murphy (1992), Acemoglu (1999;2002), and Card and Lemieux (2003). Structural work on the impact of immigration on the wages of natives include Borjas (2003), Card (2009) and Ottaviano and Peri (2011). In their analyses, these studies use the standard assumption of perfectly competitive labor markets to arrive at their econometric equation derived from profit maximizing conditions.

Here, we show the consequences of labor market imperfections on the estimation of production functions. Our main finding is that data from an economy with labor market imperfections and a given technology will be rationalized by a different technology under the assumption of competitive labor markets. Specifically, we prove that as long as the elasticities of supply to the firm are finite and not equal to one another, estimates of the production function will not be consistent. To the best of our knowledge, no prior work has addressed this misspecification in the context of labor market frictions. The only work that we are aware of that raises a similar issue is Oaxaca and Lup-Tik (2004), who note that differences in wages between two groups that are not attributable to differences in productivity may bias a nested CES estimation. They use the unexplained portion of the wage gap from a Oaxaca decomposition in a control function approach.

Simulating the estimation of the parameters of a CES production function, we show that constant and unequal elasticities of supply lead to biased estimates of the efficiency parameters. Elasticities that are a function of the wage lead to biased estimates of the elasticity of substitution between two factors. We find that under reasonable assumptions the bias can be significant enough that a true elasticity of substitution between high and low skilled labor of 2 can be incorrectly estimated as 1.4.

¹Studies measuring these parameters include Ransom and Oaxaca (2005, 2010), Ransom and Simms (2009), Falch (2010), Hotchkiss and Quispe (2009), Hirsch et al (2010) and Depew and Sorensen (2011).

2 Model

2.1 A generic non identification result

Suppose that the aggregate output of an economy is given as $Y^t = F(L_1^t, L_2^t)$ where L_1^t and L_2^t are factors of production of respective price $w_j(L_j^t)$ for $j = 1, 2$ and the production function F captures technology in use.² Data from this economy is superscripted by t representing an observational unit.³ For the sake of simplicity, we assume that the product market is perfectly competitive and use output price as a numeraire to factor prices. In contrast, the markets for factors of production may not be perfectly competitive. To formally capture departures from perfect competition, we define upward sloping supply functions linking factor prices to factors quantities, i.e. $w_j(L_j^t)$ with $\frac{\partial w_j(L_j^t)}{\partial L_j} \geq 0$. Note that the inverse elasticity of supply of factor j , i.e. $\frac{1}{\varepsilon_j(L_j^t)} = \frac{\partial \ln w_j(L_j^t)}{\partial \ln L_j} \geq 0$, is equal to 0 when $\frac{\partial w_j(L_j^t)}{\partial L_j} = 0$. In this case, factor market j is said to be perfectly competitive. Given this setting, we define an economy as a tuple $\left\{ \frac{1}{\varepsilon_1(\cdot)}, \frac{1}{\varepsilon_2(\cdot)}, F(\cdot, \cdot) \right\}$. The firms' (static) problem is to choose (L_1^t, L_2^t) to maximize profits $F(L_1^t, L_2^t) - w_1(L_1^t)L_1^t - w_2(L_2^t)L_2^t$. The FOCs to this problem are:

$$w_j(L_j^t) \left(1 + \frac{1}{\varepsilon_j(L_j^t)} \right) = \frac{\partial F(L_1^t, L_2^t)}{\partial L_j} \text{ for } j = 1, 2 \quad (1)$$

Taking logs and subtracting the two FOCs obtains:

$$\ln \frac{w_1(L_1^t)}{w_2(L_2^t)} = \ln \frac{\partial \tilde{F}(L_1^t, L_2^t)}{\partial L_1} - \ln \frac{\partial \tilde{F}(L_1^t, L_2^t)}{\partial L_2} \quad (2)$$

where $\frac{\partial \tilde{F}(L_1^t, L_2^t)}{\partial L_j} \equiv \frac{\frac{\partial F(L_1^t, L_2^t)}{\partial L_j}}{1 + \frac{1}{\varepsilon_j(L_j^t)}}$. Each economy t generates the tuple $\{w_1^t/w_2^t, L_1^t, L_2^t\}$

as data that consists of the ratio of the two factor prices $w_j^t (= w_j(L_j^t))$ for $j = 1, 2$ and two factor quantities L_j^t for $j = 1, 2$. Equation 2 is the fundamental equation of this paper. Using this equation, one can easily prove the following (non) identification theorem.

Theorem 1 *To each economy $\left\{ \frac{1}{\varepsilon_1(\cdot)}, \frac{1}{\varepsilon_2(\cdot)}, F^a(\cdot, \cdot) \right\}$ corresponds an economy $\{0, 0, F^b(\cdot, \cdot)\}$, where F^b is defined such that $\frac{\partial F^b(L_1^t, L_2^t)}{\partial L_j} = \frac{\partial \tilde{F}^a(L_1^t, L_2^t)}{\partial L_j} \equiv \frac{\frac{\partial F^a(L_1^t, L_2^t)}{\partial L_j}}{1 + \frac{1}{\varepsilon_j(L_j^t)}}$, that both generate the same data $\{w_1^t/w_2^t, L_1^t, L_2^t\}$.*

²We abstract from technological change by assuming that F is constant over observational units. This assumption is without loss of generality with respect to our key point. Similarly, we assume that the shape of the supply to the firm of factors of production is constant over time without loss of generality and for the sake of notational simplicity.

³Superscript t may represent time, space or variation between other cross-sectional units such as cohorts.

Proof. Let $\{w_1^t/w_2^t, L_1^t, L_2^t\}$ be the data generated by the economy $\left\{\frac{1}{\varepsilon_1(\cdot)}, \frac{1}{\varepsilon_2(\cdot)}, F^a(\cdot, \cdot)\right\}$.

One directly sees that there always exists an economy $\{0, 0, F^b(\cdot, \cdot)\}$ with $\frac{\partial F^b(L_1^t, L_2^t)}{\partial L_j} = \frac{\partial \tilde{F}^a(L_1^t, L_2^t)}{\partial L_j}$ such that from Equation 2 at factors quantities L_1^t and L_2^t , this economy produces relative prices equal to w_1^t/w_2^t . The two economies produce the same data $\{w_1^t/w_2^t, L_1^t, L_2^t\}$.

Theorem 1 implies that the function of production $F(\cdot, \cdot)$, without any additional information about supply elasticities (with the data in hand), is not identified non-parametrically. This non identification result has implications for the estimation of the shape of the production function as stated in the following corollary of Theorem 1. ■

Corollary 2 *Estimates of $\frac{\partial F}{\partial L_j}$ using Equation 2 and data $\{w_1^t/w_2^t, L_1^t, L_2^t\}_{t=0}^T$ are inconsistent when the two factors have different elasticities of supply, i.e. $\varepsilon_1(\cdot) \neq \varepsilon_2(\cdot)$.*

Proof. Let the true economy be $\left\{\frac{1}{\varepsilon_1(\cdot)}, \frac{1}{\varepsilon_2(\cdot)}, F^a(\cdot, \cdot)\right\}$. The associated generated data are $\{w_1^t/w_2^t, L_1^t, L_2^t\}_{t=0}^T$, i.e. a sample of T observations of this economy.⁴ Assuming the type of production is known but the parameters not, using these data one estimates $\ln \frac{\partial \tilde{F}(L_1, L_2)}{\partial L_1} - \ln \frac{\partial \tilde{F}(L_1, L_2)}{\partial L_2} = \ln \frac{\partial F(L_1, L_2)}{\partial L_1} - \ln \frac{\partial F(L_1, L_2)}{\partial L_2} + \ln \left(\frac{1 + \frac{1}{\varepsilon_2(L_2)}}{1 + \frac{1}{\varepsilon_1(L_1)}} \right)$. The last term only collapses when $\varepsilon_2(L_2) = \varepsilon_1(L_1)$, a special case of which is met when both factor markets are perfectly competitive. The sign of the bias in the ratio of marginal productivity will be given by the sign of $\ln \left(\frac{1 + \frac{1}{\varepsilon_2(L_2)}}{1 + \frac{1}{\varepsilon_1(L_1)}} \right)$. ■

2.2 CES production functions

Let technology be summarized by a CES production function, i.e. $F(L_1, L_2) = [\lambda L_1^\rho + (1 - \lambda) L_2^\rho]^{1/\rho}$ with $\rho \in (-\infty, 1]$. The parameter λ measures the efficiency of factor 1 relative to factor 2 whereas the parameter ρ determines σ the elasticity of substitution between factors, i.e. $\sigma = \frac{1}{1-\rho}$. Given the CES technology, Equation 2 reads as:

$$\ln \frac{w_1(L_1^t)}{w_2(L_2^t)} = \ln \frac{\lambda}{1-\lambda} + (\rho - 1) \ln \frac{L_1}{L_2} - \ln \left(\frac{1 + \frac{1}{\varepsilon_1(L_1)}}{1 + \frac{1}{\varepsilon_2(L_2)}} \right) \quad (3)$$

It is clear that as long as the last term does not collapse, an estimate $\hat{\lambda}$ of λ will

⁴The data may vary across observational units for various reasons. An interesting interpretation would be that the wage curves $w_1^t(L_1^t)/w_2^t(L_2^t)$ differ by a scalar factor across units, i.e. $w_1^t(L_1^t)/w_2^t(L_2^t) = \mu^t w_1(L_1^t)/w_2(L_2^t)$. The equilibrium data are then different across observational units but the elasticities of supply remain similar. Putting $\ln \mu^t$ on the right hand side of Equation 2 produces an "error" term that is generally added to the equation in empirical applications.

be biased (upward or downward). It is also clear that as soon as the elasticities in the last term vary with L_j^t , and hence $\ln L_j^t$, an estimate $\hat{\rho}$ of ρ will be biased (again upward or downward). Two particularly interesting cases arise.

Case 1: *factors have constant supply elasticities, i.e. $\varepsilon_j(L_j) = \varepsilon_j$.* In this case,

Equation 3 becomes:

$$\ln \frac{w_1(L_1^t)}{w_2(L_2^t)} = \ln \frac{\hat{\lambda}}{1 - \hat{\lambda}} + (\rho - 1) \ln \frac{L_1}{L_2} \quad (4)$$

where $\ln \frac{\hat{\lambda}}{1 - \hat{\lambda}} = \ln \frac{\lambda}{1 - \lambda} - \ln \left(\frac{1 + \frac{1}{\varepsilon_1}}{1 + \frac{1}{\varepsilon_2}} \right)$. The estimate $\hat{\lambda}$ of λ will suffer an upward

bias when $\frac{\varepsilon_2}{\varepsilon_1} < 1$ and a downward bias otherwise. However, the estimate $\hat{\rho}$ will be consistent in this case.

Case 2: *supply elasticities are such that $1 + \frac{1}{\varepsilon_j(L_j)} = \alpha_j L_j^\beta$, with $\beta \in \mathbb{R}$ and $\alpha_j > \frac{1}{L_j^\beta} > 0$.*⁵ In this case, Equation 3 becomes:

$$\ln \frac{w_1(L_1^t)}{w_2(L_2^t)} = \ln \frac{\hat{\lambda}}{1 - \hat{\lambda}} + (\hat{\rho} - 1) \ln \frac{L_1}{L_2} \quad (5)$$

where $\ln \frac{\hat{\lambda}}{1 - \hat{\lambda}} = \ln \frac{\lambda}{1 - \lambda} - \ln \left(\frac{\alpha_1}{\alpha_2} \right)$ and $\hat{\rho} = \rho - \beta$. Case 1 is clearly a special case

met when $\beta = 0$ and where $\alpha_j = 1 + \frac{1}{\varepsilon_j}$. The bias in the estimate $\hat{\lambda}$ of λ is therefore similar to that discussed in Case 1. In contrast to Case 1 though, as soon as $\beta \neq 0$, an estimate $\hat{\rho}$ of ρ will suffer a bias: downward bias when $\beta > 0$ and upward when $\beta < 0$.

Theorem 1 is illustrated in Figure 1 using the specification of Case 2. The marginal costs and marginal productivity of factor 1 relative to factor 2 are plotted on two panels. The left panel corresponds to an economy $\{0, 0, \tilde{F}(\cdot, \cdot)\}$, i.e. $\beta = 0$, $\alpha_j = 1$, whereas in the right panel corresponds to an economy $\left\{ \frac{1}{\varepsilon_1(\cdot)}, \frac{1}{\varepsilon_2(\cdot)}, F(\cdot, \cdot) \right\}$.⁶ On each panel, the equilibrium condition captured in Equation 2 corresponds to the crossing point of the marginal cost curves with the marginal productivity curve. In each panel we consider three supply curves capturing shifts in relative supply over time or across countries. The observed

⁵Note that solving for the supply function $\ln w_j(L_j)$ given our definition $1 + \frac{1}{\varepsilon_j(L_j)} = \alpha_j L_j^\beta$ yields: $\ln w_j(L_j) = \frac{\alpha_j}{\beta} L_j^\beta - \ln L_j + \text{const}_j$ where const_j is a constant of integration. Re-write $\frac{\alpha_j}{\beta} L_j^\beta$ as $\alpha_j \left(\frac{L_j^\beta - 1}{\beta} + \frac{1}{\beta} \right)$ and note that $\lim_{\beta \rightarrow 0} \frac{L_j^\beta - 1}{\beta} \rightarrow \ln L_j$ (see the Box-Cox transformation). Let $\text{const}_j = w_j - \frac{1}{\beta}$, $\alpha_j = 1$ and $\beta \rightarrow 0$. We then have $\ln w_j(L_j) \rightarrow w_j$ which corresponds to the perfect competition case.

⁶In these simulations, we used a CES production function and the parametric shape discussed in section to calibrate the supply curves.

data are represented by the three blue points $\left(\frac{L_1^t}{L_2^t}, \frac{w_1^t}{w_2^t}\right)$. As stated in Theorem 1 these points are identical in both panels. Note that, in contrast to perfect competition (the left panel), under monopsony (the right panel) these data do not trace out a demand curve consistent with the true production function.

3 Monte Carlo Simulation

The simulation exercise considers an economy with CES technology $F(L_1, L_2) = [\lambda L_1^\rho + (1 - \lambda)L_2^\rho]^{1/\rho}$ and supply functions as in Case 2 above, i.e. $\frac{1}{\varepsilon_j(L_j)} = \alpha_j L_j^\beta - 1$. This economy generates data $\{w_1^t/w_2^t, L_1^t, L_2^t\}_{t=0}^N$. The data generating process is then:

$$\ln \frac{w_1^t}{w_2^t} = \ln \frac{\hat{\lambda}}{1 - \hat{\lambda}} + (\hat{\rho} - 1) \ln \frac{L_1^t}{L_2^t} + u^t \quad (6)$$

where $\ln \frac{\hat{\lambda}}{1 - \hat{\lambda}} = \ln \frac{\lambda}{1 - \lambda} - \ln \left(\frac{\alpha_1}{\alpha_2}\right)$ and $\hat{\rho} = \rho - \beta$. The random term u^t is independently distributed and follows a Normal distribution⁷ $N(0, 0.2^2)$ whose variance is scaled such that the R^2 of our regressions is about 0.85. Factor quantities L_j^t are independently and uniformly distributed over $[5, 15]$.

We consider four economies that share the same technology with elasticity of substitution equal to 2, i.e. $\rho = 0.5$, and equal efficiencies for both factors. $\lambda = 0.5$. However, these four economies face different factor market structures. In particular:

- (Economy 1) $\alpha_1 = 0.3$, $\alpha_2 = 0.1$ and $\beta = \frac{1}{4}$,
- (Economy 2) $\alpha_1 = \alpha_2 = 0.3$, and $\beta = \frac{1}{4}$,
- (Economy 3) $\alpha_1 = 1.5$, $\alpha_2 = 2$, $\beta = 0$ and,
- (Economy 4) $\alpha_1 = \alpha_2 = 1$ and, $\beta = 0$.

Economy 1 corresponds to our benchmark economy with supply elasticities of 2.14 and 1.22 for factor 1 and 2 respectively. Economy 4 is the polar economy that faces perfectly competitive factor markets. Two intermediate cases are also examined. In economy 2, factor markets are imperfect but the elasticities of supply are equal for both factors. In economy 3, the elasticities of supply are constant though different for the two factors. For each economy, we generate 1,000 samples of size $N = 100$. For each sample we estimate the parameters $\hat{l} = \ln \frac{\hat{\lambda}}{1 - \hat{\lambda}}$ and $\hat{\sigma} = \frac{1}{1 - \hat{\rho}}$ of Equation 6 using Ordinary Least Square technique. The simulation results are reported in Table 1. As expected, the estimates of both \hat{l} and $\hat{\rho}$ are biased in Economy 1. The relative efficiency of factor 2 is over estimated and the magnitude of the bias is equal to $\ln \left(\frac{0.3}{0.1}\right) = 1.10$. The estimate of the elasticity of substitution is also biased, i.e. $\hat{\sigma} = \frac{1}{1 - \hat{\rho}} = \frac{4}{3} \approx 1.33$.

⁷An interpretation of u is provided in footnote 4.

It is striking to see that given our parametric setting, at feasible magnitudes of supply elasticities⁸, one estimates the elasticity of substitution to be 1.33 (as Katz and Murphy for instance) whereas the true elasticity is equal to 2.

Finally, as expected, estimates of the relative efficiencies are biased in Economy 3 but unbiased in Economy 2 while estimates of the elasticity of substitution are unbiased in Economy 3 but biased in Economy 2.

4 Conclusion

Recent work has established strong evidence that the labor elasticity of supply to the firm is finite. Production function estimation based on a representative firm model had typically assumed that the firm faces perfectly competitive input markets. Finite elasticities imply that the firm faces profit maximizing conditions that differ from the standard model. Here, we have explored the economic consequences of this assumption being violated.

Our first finding is that equilibrium data on input factors (wages and quantities) and an incorrect assumption of perfectly elastic labor supply to the firm will be rationalized only with a production function that differs from the true production function in the economy. This is demonstrated in a clear example shown in Figure 1.

We then explore this finding in the context of a CES production function. Elasticities that vary with the wage will generate biased estimates of the elasticity of substitution as well as the efficiency parameters. Even constant but unequal elasticities will bias estimates of the efficiency parameters, though not the elasticity of substitution. It should be noted that, in the context of a Nested-CES, incorrectly estimating the efficient parameters in the bottom level of the nest may cause higher level elasticities of substitution to be inconsistently estimated. A Monte Carlo simulation shows that, under reasonable parameter values, the bias in the elasticity of substitution between two groups, say high and low skilled labor, may be significant. Specifically, we show that a true elasticity of 2 may be estimated as 1.4.

These findings provide a caveat for production function estimation. Ideally, a structural model would use data that would also allow for estimation of the labor elasticity of supply to the firm in order to correctly specify the production function estimation model.⁹ However, such firm level data is hard to come by.

A potentially valuable method for applied researchers to follow is the control function approach used by Lup-Tick and Oaxaca (2004) in their paper on gender biased technological change, where they consider that male and female wages may differ for reasons other than differences in marginal productivity,

⁸This numbers are comparable with estimates from the studies listed in Footnote 1

⁹Dupuy and Marey (2008) estimate Equation 5 for skilled and unskilled labor in the US between 1963 and 2005. They find that $\hat{\sigma} = \frac{1}{1-\hat{\rho}}$ shifts twice over time. Although Dupuy and Marey (2008) argue in favor of technical change as an explanation thereof, Corollary 2 indicates that changes in the relative supply elasticities of skilled and unskilled labor could have generated these shifts.

such as taste based discrimination. In a more general context, their approach of controlling for the unexplained wage gaps from a Oaxaca decomposition may also help tackle the issues raised here, where different labor elasticities of supply to the firm affect wages along with differing marginal productivity.

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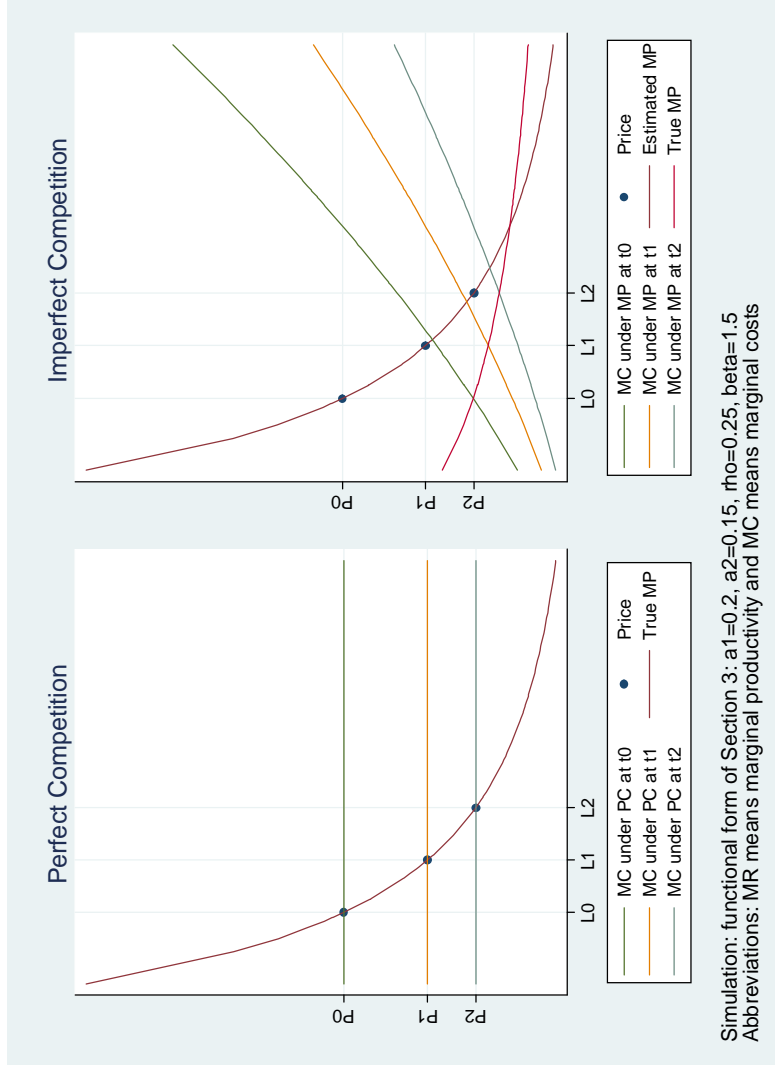


Figure 1: Illustration of Theorem 1. On the left panel the input markets are perfectly competitive whereas on the right panel supply of inputs are less than perfectly elastic. Nevertheless, due to differences in technology across economies, both economies generate the same data (blue dots).

Table 1: Monte Carlo Simulations on four economies.

	Economy 1		Economy 2		Economy 3		Economy 4	
	Mean	95% CI	Mean	95% CI	Mean	95% CI	Mean	95% CI
$\ln \frac{\lambda}{1-\lambda}$	0.00		0.00		0.00		0.00	
TRUE	-1.10	[-1.14;-1.06]	0.00	[-0.04;0.04]	0.29	[0.25;0.33]	0.00	[-0.04;0.04]
Estimated	-1.10		0.00		0.29		0.00	
True bias								
σ	2.00		2.00		2.00		2.00	
TRUE	1.34	[1.18;1.50]	1.34	[1.18;1.50]	2.00	[1.62;2.38]	2.00	[1.62;2.38]
Estimated	-0.66		-0.66		0.00		0.00	
True bias								