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Arbeitspapiere/ Working Papers

Aniela Wirz

To my Wife, with Love!

Does Within-household Specialisation Explain
Husbands' Better Job-education-match?

To my wife, with love!*

Does within-household specialisation explain
husbands' better job-education-match?

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Abstract

Married male workers are found to have a lower incidence of over-education. A theoretical explanation for this phenomenon is lacking. We test in our study whether the traditional specialisation of spouses' time between home and market production tends to improve a husband's job-education-match (JEM). We test this hypothesis first by drawing on the method used in the marriage wage premia literature based mainly on the model of Becker (1985). In addition, we perform a new test following the theory of François (1998), which requires less restrictive assumptions. Overall, our results show that within-household specialisation (WHS) explains a substantial part of the superior JEM of husbands, regardless of whether a wife's labour market participation (experience) or both spouses housework hours are used to measure specialisation. The results and in particular the independent and significant impact of women's housework hours on their husbands' JEM, however, speak clearly in favour of François' theory and against the explanation of Becker. Testing for an endogeneity bias due to a possible sorting process of more able husbands with "traditional" spouses or a measurement error of the JEM does not alter these conclusions.

Keywords: Overeducation; Household models, Human capital, Labour productivity

JEL classification: I21; J16; J24;

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Saying:

“Behind every great man there is a great woman”

1 Introduction

The above saying reflects the common idea, that a husband’s success is partly due to the support of his wife. Married men are indeed found to have a significantly higher probability of working in a job that matches their educational qualifications and hence a lower probability of working overeducated than single men. A theoretical explanation for this better *job-education-match* (JEM) of husbands is lacking so far in the overeducation literature.¹ Recent research in this area emphasises that workers having educational qualifications superior to the requirements of their job simply lack other productive characteristics (ability, motivation, effort, etc.). Hence, they have lesser chances of getting an appropriate job.² The question then arises whether and how being married affects such unobservable productive characteristics thereby improving the JEM of men.

We aim to test whether “traditional” specialisation of married men and women between market and home production explains the improved JEM of husbands. Or, asked differently: “Do married men get better jobs if they have a ‘stay-at-home’ spouse doing the housework for them?” If this is the case, “What enables married men to get a better job given their education?”

We do this by testing two theoretical explanations. The first one is Becker’s (1985) seminal work on specialisation generally used to explain higher wages of married men (marriage wage premia). Basically, Becker stipulates that married men are more productive in paid work because they are able to channel all their energy and effort into paid work when their wives do the housework for them. Although this model is intuitively quite appealing and therefore very much in line with the introductory saying, recent empirical research seriously questions it. Using more detailed information about housework hours of men Hersch and Stratton (2000) could not confirm the causal mechanism formulated in the model. This might be due to the very restrictive assumptions underlying it. This becomes obvious in the fact that *within-household specialisation* (WHS) is assumed to have a positive impact on the husband’s productivity exclusively by lowering his housework burden when getting married. This model excludes many cases, where WHS is increased and this burden is not lowered. For instance, when having children, husbands may frequently be found to increase their amount of housework, e.g. from 7 to 15 hours a week, while their wives increase their housework hours from 15 to 50. Hence, in these cases the degree

¹The term bad JEM is used to describe a worker being “overeducated” for his job throughout the paper. Similarly a good JEM describes a worker being “adequately educated” in the terminology of the overeducation literature. Furthermore, the terms husbands and partners are used as mutual equivalents.

²See Bauer (2002) and Chevalier (2003). For an overview on the overeducation literature see Sloane (2002), Hartog (2000) and Groot and Maassen van den Brink (2000).

of WHS increases strongly, but the husband's productivity is assumed to be unaffected.

Therefore, in a next step, we develop a new test of the specialisation hypothesis, following the theory of within-household interaction of François (1998). Within his theoretical framework WHS has an impact on a husband's work effort and hence his JEM in a more general way. The assumed behaviour of spouses is suggested to be the result of a bargaining process between two distinct persons with differing interests. In addition, leisure and housework hours are clearly distinguished contrary to Becker's approach. In short, François explains that the work effort of a husband is higher than the work effort of a single man, because having a good job not only allows him to earn a high wage but in addition improves his bargaining position within the household. Thus his work motivation is increased by his gains from within-household trade. This trade gain is the amount of home production taken over by his wife, which he loses when he shirks and gets fired. This model thus argues on a less romantic level than the introductory saying assuming implicitly that it is the support of a woman which helps a man to be more productive. Strongly simplified, it suggests that a husband's fear of having to do a larger share of the home production increases his motivation to get and keep a good job. It is important to note, that in this case the amount of housework hours done by the wife has an effect on the work effort and JEM of her husband independent of his's own housework hours.

Thus, the contribution of this study to the literature is twofold. First, we try to explain whether WHS explains the improved JEM of married men. Does "the women behind the man" matter or does only his own human capital matter for his productivity at market work. Second, we perform a new test for the specialisation theory in order to explain more precisely why specialisation has such an impact. We can do this because information on both spouses market work hours and housework hours are available in our data set.

More generally, this paper contributes to a better understanding of gender differences in the labour market by choosing an approach opposite to the conventional one, which focuses on women's situation in the labour market. The idea is that if we know what contributes to the success of men at market work, we should also know better what makes women worse off in the labour market compared to men.

In our empirical analysis we use the Swiss Household Panel (SHP) data for 1999. Using probit analysis we test whether husbands with a wife being strongly specialised in home production have on average a significantly higher probability to have a good JEM than husbands with a wife sharing the household work more equally. This degree of specialisation (WHS) will be measured in several ways. In a first step, we take women's market work (weekly market work hours, experience in years) as a proxy for WHS following a common method in the marriage wage premia literature. In a second step, we use directly both spouses' home production (weekly hours of housework) to measure the degree of WHS. This allows us also to distinguish between the two theoretical models. If specialisation has an effect mainly through the husband's lower home production,

then this supports Becker’s model (1985). On the contrary, if the husband’s higher probability to obtain a good JEM is explained less by his own housework but rather by his wife’s housework hours and / or the difference between their housework hours, we conclude that François’ (1998) model applies.

A particular emphasis will be given to measure an eventual impact of unobserved heterogeneity of husbands respectively couples. More precisely, we adapt our model to test whether a sorting process of more able husbands with “traditional” spouses is at work instead of the expected specialisation effect. In addition, we test whether measurement errors of the JEM could possibly bias the results.

The paper is organised as follows. Section 2 gives a more detailed discussion of the relevant literature. In Section 3, the methodology and model specification issues are detailed. In Section 4, the data used is described and in Section 5 the descriptive statistical evidence presented. Section 6, shows the full model estimation results, and Section 7 concludes.

2 Theoretical and empirical background

There are two strands of literature related to our study. The first one is the overeducation literature, which provides no theory for the improved *job-education-match* (JEM) of husbands but a vast array of empirical evidence about the relationship between the JEM and family commitments. The second one is the research about *within-household specialisation* (WHS) or bargaining and its impact on a husband’s work effort. This literature covers both theoretical and empirical approaches. As these two strands of literature were not related so far they are treated separately in the Subsections 2.1 and 2.2. The theories of Becker (1985) and François (1998) are discussed in-depth in Subsection 2.2.

2.1 Quality of the JEM and family commitments

There is clear empirical evidence that marital status decreases significantly the probability for men to be overeducated.³ However, this result is mostly contingent on a gender-specific test of the determinants of overeducation, owing to the fact that family responsibilities have frequently an adverse impact on women’s overeducation incidence. As an exception, Büchel and Battu (2002) do not find a significant impact of marriage for men in Germany even if detailed interaction terms are used. Differences in the model specifications are, however, important.⁴ Finally, Green and MacIntosh (2002), using data from the UK, find a significant

³See Van der Meer and Batenburg (2002) for the Netherlands, Sloane, Battu and Seaman (1996) for the UK and Wirz and Atukeren (2004) for Switzerland.

⁴Differences in specification concern mostly variables approximating unmeasured productive differences of workers, the location of the household and controls for part-time working. Büchel and Battu (2002) add controls for the location of the household, add a variable measuring the importance of occupational success but they do not control for part-time working of either gender.

negative impact of marital status on the probability of working over-skilled even when gender is not interacted with.

Hence, a positive impact of marital status on the quality of a husband's JEM is rather the norm. A theoretical explanation for this better JEM or lower overeducation incidence of married men is lacking so far in the overeducation literature.

The only theoretical model explaining a possible relationship between geographical factors, family commitments and overeducation is given by the theory of "differential overqualification" following the seminal paper of Frank (1978). According to this theory, married women have a higher risk of working in jobs for which they are overqualified. They are supposed to be "tied movers" or "tied stayers" in the sense that their job search is undertaken under the condition that the husband's job search is optimised first. If the size of the local labour market, where their husband found a job, is small, their job search is restricted and they have an increased incidence of overeducation. Following this theory, there is no reason for husbands to have a better JEM than single men. So far, empirical evidence for this theory is mixed as shown by Büchel and Battu (2002) for Germany, and McGoldrick and Robst (1996) for the US. On the contrary, Dolton and Silles (2001) find empirical support for rigidities, linked to geographical factors and family commitments to affect the matching process of both gender. The willingness to relocate for the first job and having children prior to the first job improve the chances of having a good JEM respectively lower the chances of working overeducated for young graduates in the UK significantly. The impact of spatial factors is clearly confirmed by Büchel and van Ham (2003).

Overall, these results suggest that motivation and effort to get a good job, for instance by looking further than the local labour market, seem to be an important part of the unmeasured individual characteristics determining the quality of the JEM. Similarly, marital status and family commitments seem to affect the efficiency of this matching process. The causal mechanisms governing these effects, however, are quite unclear so far.

2.2 WHS and the work effort of husbands

The theory of specialisation, which will be tested to explain husbands' improved JEM, goes back to Becker (1985, 1991). In his theory of the family (1991) he stipulates that it is efficient for spouses to specialise between home and market production according to their respective comparative advantages and preferences. Therefore, a woman with a lower wage rate and / or higher productivity in household production than her husband tends to lower her market work hours and increase her home production. Such WHS is assumed to have a positive effect on a husband's productivity in market work.

How such positive specialisation effects materialise remains open. Becker (1985) argues that total effort is limited and thus any effort allocated to home production reduces the effort available for market work. Hence, married men may benefit from specialisation. By reducing their housework hours they can

concentrate their time and effort on their market production and therefore improve their chances of getting a better job. However, empirical tests of this theory, mostly done while analysing marriage wage premia, showed ambiguous results. Women’s lower employment status, for example, seemed clearly to be correlated with a higher work effort of their husband as measured by higher wages [see Hersch and Stratton (2000) for an overview]. Surprisingly, this higher wage could not be explained by lower housework hours of the husband, as the theory stipulates. Becker’s specialisation explanation was rejected by Hersch and Stratton (2000) on these grounds. Bonke, Datta Gupta, and Smith’s (2003) results from a Danish data set also hint at a more complex relationship between housework hours and a husband’s wage.

More recent models of within-household interaction are based on less restrictive assumptions than the model of Becker. Following the seminal work of Manser and Brown (1980) and McElroy and Horney (1981, 1990), a wide array of bargaining models explain the allocation of consumption and production between spouses. In these models the utility functions of spouses are allowed to differ. The time allocation decision process can be determined by the respective bargaining power of spouses mostly approximated by their respective earning power. If home production is taken as an “undesired activity”, then the spouse with a higher bargaining power will contribute less to home and more to market production. However, the question whether and how specialisation between spouses may benefit the husband’s success in the labour market (improved JEM) remains unanswered by most of these models.

To our knowledge, François (1998) formulated the only model allowing for a bargaining process between spouses about their respective time allocation, thereby explaining why men should have a higher probability of having a good job if they are married.⁵ Based on the assumption that household production can only partly be substituted by market production and that spouses differ in employment characteristics on average, he argues that it is reasonable to assume that one spouse specialises in home production. Married men are then expected to outperform single men since they are able to trade with their spouses to undertake the provision of household services. So far, the argumentation is not totally different from Becker’s explanation of the specialisation behaviour within households. But the mechanism assumed to improve a husband’s productivity is quite different. In François’s model the labour market entails efficiency wages and thus he stipulates that a husband is less likely to shirk than single men, because of the higher costs of getting fired. He not only loses his efficiency wage income but also his gains from within-household trade.⁶ Thus, a husband’s

⁵Skatun (2004) recently formulated an alternative bargaining model explaining the male wage premium by a better bargaining position of married men when compared to single men. The higher is the earnings potential of the female partner the higher is the expected wage of the husband. However, he does not make any assumptions about married man being indeed more motivated and productive at work.

⁶François (1998) makes many assumptions to simplify the analysis although they are not critical. For instance, all results of the analysis hold also if workers are simply assumed to be less likely to shirk when their spouse has a bad job. Hence, this model does not exclude other models of within-household interactions based on altruistic motives of partners for instance.

effort at work is expressed as a function of the amount of home production he gets from his wife (benefit in kind) in exchange for his higher wage income. This is a more general mechanism than the one suggested by Becker, where a husband's effort at work is increased exclusively by the amount of housework hours the husband does not have to do compared to a single man. Following this logic, an increase of women's total housework, e.g. after the birth of a child, is not thought to have an independent impact in the model of Becker. On the contrary, it should have an impact following the theory of François. However, the mechanism of within-household interaction stipulated by François (1998) was never tested micro-econometrically.

Hence, so far it is still unclear what factors contribute to a better chance of husbands to have a good job given their educational level and whether specialisation of spouses' time between home and market production indeed has a positive impact on husbands' work motivation and effort. This is the focus of our study.

3 Empirical approach

For the purpose of our analysis we create a variable "good JEM" denoted GJEM. GJEM is a binary variable, taking value 1 if the husband's qualifications are adequate for his job and 0 if his qualifications are superior to his job. In doing this we invert the notation generally used in the overeducation literature, where the overeducation variable takes value 1 when a worker is overeducated and 0 otherwise. Apart from that we follow the basic model used in the overeducation literature in a first step (*Model 1*). The probability of a husband having a GJEM of 1, denoted by variable y^h , is then formulated as an index model, where y^{h*} is an underlying latent variable expressed as a function of the productive characteristics of the husband such as experience or education summarized in vector x^h and the degree of within-household specialisation (WHS) represented by variable s .

For estimation we use observations of N households (or couples), which are assumed to be independent and identically distributed following *Model 1*. The household as the unit of observation is denoted by i . Hence, the relationship to be estimated is the following,

$$y_i^h = \begin{cases} 1 & \text{if } y_i^{h*} > 0 \\ 0 & \text{else} \end{cases}$$

$$y_i^{h*} = \beta_1 + x_{2i}^h \beta_2 + x_{3i}^h \beta_3 + \dots + x_{Li}^h \beta_L + s_i \beta_s + \varepsilon_i \quad (1)$$

where β_1 denotes the constant, $\beta_2, \beta_3, \dots, \beta_L$, the coefficients of the productive characteristics of the husband. It is finally β_s , the coefficient of variable

In such a model a husband then increases his work effort mainly because he wants his family to have the highest possible utility level.

s , we are interested in, as it measures the impact of WHS on the probability of a husband to have a good job-education-match. The error term of the equation, ε_i , is a continuously distributed random variable (i.i.d.), independent of the explanatory variables x_{li}^h , $l = 1, 2, 3, \dots, L$, and s_i . Assuming a standard normal distribution of ε_i , we estimate this relationship by means of a probit model.

We will test three variations of this basic *Model 1* denoted by a), b) and c), where in each case equation (1) is slightly modified. In a first step, WHS will be approximated by a wife's labour market participation variables, which are market work hours and experience in years following the method mostly used in the analysis of marriage wage premia. This variation is simply equation (1a) where s_i is in fact s_i^w as WHS is represented by the wife's characteristics only. The impact of WHS is again measured by β_s the coefficient of the variable s_i^w .

$$y_i^{h*} = \beta_1 + x_{2i}^h \beta_2 + x_{3i}^h \beta_3 + \dots + x_{Li}^h \beta_L + s_i^w \beta_s + \varepsilon_i \quad (1a)$$

In a second and third step, WHS will be approximated directly by housework hours of both spouses s_i^w and s_i^h . In equation (1b) we use this information individually and in equation (1c) we use the in-between difference of these housework hours ds_i^{wh} as measures for WHS, data which is available in the SHP (1999).

$$y_i^{h*} = \beta_1 + x_{2i}^h \beta_2 + x_{3i}^h \beta_3 + \dots + x_{Li}^h \beta_L + s_i^w \beta_{sw} + s_i^h \beta_{sh} + \varepsilon_i \quad (1b)$$

$$y_i^{h*} = \beta_1 + x_{2i}^h \beta_2 + x_{3i}^h \beta_3 + \dots + x_{Li}^h \beta_L + ds_i^{wh} \beta_{ds} + \varepsilon_i \quad (1c)$$

These last two variations provide us first an alternative and more direct measure of WHS and its impact on GJEM ($\beta_{sw}, \beta_{sh}, \beta_{ds}$). Second, these variations allow us to investigate more precisely how WHS has an impact on GJEM and hence to discriminate between the two theoretical models. If specialisation takes its effect mainly through the husband's lower home production, as measured by the size and statistical significance of coefficient β_{sh} in equation (1b), then this lends support to Becker's (1985) more restrictive model of specialisation effects. On the contrary, if the husband's higher probability to get a good JEM is explained less by his own housework (β_{sh}) but rather by his wife's housework hours (β_{sw}) in equation (1b) and / or the difference between their housework hours (β_{ds}) in equation (1c), we conclude that rather François' model (1998) applies, as discussed in detail in Section 2.2.

3.1 Neglected heterogeneity of households and measurement problems of quality of the JEM

There are two main problems arising in this estimation setting. The first is unmeasured heterogeneity of couples with respect to important characteristics such as a husband's ability and his wife's preference for housework. The second

eventual problem is measurement error of the JEM. Such omitted variables or measurement error could bias the results if they are correlated with the explanatory variable of interest WHS respectively s_i^w , s_i^h or ds_i^{wh} . One of the main critical assumptions in *Model 1* is then violated and the simple probit estimation will not give consistent results.

Omitting a variable which controls for the couples' important characteristic could indeed bias the results, if there is a sorting process, whereby husbands which are more able and motivated to get a good job more frequently get married to "traditional" women specialising in home production. The positive correlation between WHS and the propensity of a husband to have a good JEM ($GJEM = 1$) then merely reflects this sorting mechanism or unmeasured heterogeneity of households and not a causal relationship as stipulated by the theories about specialisation effects.

The argumentation is the same in the presence of measurement error problems. These could be important in our case, where the quality of the JEM of a husband is a subjectively measured variable, see Section 4 for further discussion. If such a measurement error is correlated with WHS, the results are biased. This could be the case, if for instance above average self-confident husbands qualifying themselves as having a bad JEM, because they wrongly judge themselves as being able to do a better job, are systematically more often married to "traditional" women. The negative impact of WHS on the quality of the JEM measured in *Model 1* would then merely reflect this sorting and measurement error effect.

In these cases the true model would be *Model 2*, a recursive equation system, where WHS of household i , denoted by the variable s^w , is modelled endogenously. Instead of equation (1a) the following equation system (2a) is then to be estimated,

$$s_i^w = \begin{cases} s_i^{w*} & \text{if } s_i^{w*} = \gamma_1 + x_{2i}^w \gamma_2 + x_{3i}^w \gamma_3 + \dots + x_{Ki}^w \gamma + \nu_i > 0 \\ 0 & \text{else} \end{cases} \quad (2a)$$

$$y_i^h = \begin{cases} 1 & \text{if } y_i^{h*} = \beta_1 + x_{2i}^h \beta_2 + x_{3i}^h \beta_3 + \dots + x_{Li}^h \beta_L + s_i^w \beta_s + \varepsilon_i > 0 \\ 0 & \text{else} \end{cases}$$

where x_i^w is a vector of K , $k = 1, 2, 3, \dots, K$, variables representing a wife's personal characteristics such as age, number and age of children including a first element being unity. Depending on the measure of WHS used, additional variables such as parents' education, social status and the spouses' individual attitude towards gender measures and national traditions are added. x_i^h is again a vector of L variables representing the productive characteristics of the husband. ν_i and ε_i denote the error terms of the equations, which are assumed to be joint normally distributed, each with mean zero, unit variance, and correlated due to the omitted variable or measurement error ($\rho_1 \neq 0, \rho_1 = (Corr(\nu_i, \varepsilon_i))$).

The estimation is complicated by the fact that the WHS represented above by the variable s_i^w is in most cases a left censored variable, e.g. when approximated by working hours, experience or housework hours. Exception to that is the case where specialisation is approximated by the difference between spouses housework hours (ds_i^{wh}). Thus *Model 2* is in these specifications a non-linear equation system with two (or three) limited dependent variables. Such an equation system cannot be estimated by two-step procedures, because the predicted value of s_i^w in the equation system (2a), resulting from a non-linear transformation in the first step, cannot be supposed to have a standard normal distribution. Hence, the second stage estimation cannot be a standard probit estimation, as the expectedated value of an undefined distribution cannot be computed. Estimation of the equation system by Maximum Likelihood Estimation (MLE) is then efficient although cumbersome. Given this fact, it is appropriate to test beforehand whether there is indeed reason to assume an endogeneity bias. We do this using a method suggested by Rivers and Vuong.⁷

For this endogeneity test we estimate in a first step a reduced form, where WHS represented by variable s_i^w is estimated as a function of all exogenous variables of *Model 2*. From this equation we obtain the estimated value of the residual $\widehat{\nu}_i$, which is added to the probit estimation of GJEM, denoted by the variable y_i^h in a second step.⁸

$$y_i^h = \begin{cases} 1 & \text{if } y_i^{h*} = \beta_1 + x_{2i}^h \beta_2 + \dots + x_{Li}^h \beta_L + s_i^w \beta_s + \widehat{\nu}_i \beta_{si} + \epsilon_i > 0 \\ 0 & \text{else} \end{cases} \quad (3a)$$

The probit t statistic on $\widehat{\nu}_i$ is then a valid test of the null hypothesis that s_i^w is exogenous. This conclusion is dependent on the instruments for specialisation x_{ki}^w being themselves exogenous. If the null hypothesis of exogeneity of s_i^w has to be rejected, then *Model 2* is the true model. Estimation of the non-linear equation system (2a) by MLE is then appropriate. Otherwise *Model 1* is the true model and one-step probit estimation of the husbands quality of the JEM as shown by equation (1a) applies.⁹ The test of exogeneity is, however, equally applied for equation (1c).

⁷See Wooldridge (2002, pp. 473-477).

⁸The distribution of ν has no relevance under the null hypothesis. Therefore, the test of exogeneity is valid without assuming normality or homoscedasticity of ν , and it can be applied very broadly, even if s_i^w is a limited dependent variable [see Wooldridge (2002, p. 474)].

⁹The best method to tackle the problem of unobserved heterogeneity would be the estimation of the probability of a good JEM of the husband by a fixed-effect logit model. This method has the advantage of controlling thoroughly for any individual specific unmeasured component. Unfortunately problems of measurement errors are increased by this method. In addition, the relevant sample size is strongly reduced by this method as only the individuals changing their JEM quality identify the model. Hence, a sufficiently long panel would be needed in order to get reliable results. Given the fact that only three waves of the Swiss Household Panel (SHP) were available at the time this study was undertaken, we leave the estimation with this method to subsequent research.

4 Data and sample statistics

4.1 Data

We use the Swiss Household Panel data (SHP) of 1999 in our empirical analyses. This data set includes information on approximately 7000 individuals from around 5000 households. The data are obtained by computer-assisted telephone interviews and adjusted by appropriate weights for design and non-response bias and thus the data set is representative of the permanent resident population of Switzerland.

We are interested in two main attributes of this data set. First, it contains information about the quality of the JEM of respondents and second it contains detailed informations about both spouses or partners living in a consensual partnership and in particular about their labour market and home production.

JEM: The information about the JEM is based on the self-assessment of respondents. Workers are asked how they estimate their qualifications with regard to their current job. If they say their qualifications are superior to their job they are then considered overeducated or not having a “good JEM” (GJEM=0). When a worker reports that his qualifications correspond to his job, he is not considered overeducated but having a “good JEM” (GJEM=1). Individuals reporting to have educational qualifications not sufficient or not relating to their job are excluded from the analysis in order to increase the homogeneity of the sample.

Besides this subjective measure there are two alternative methods used for measuring educational mismatch; the objective measure based on a systematic job evaluation by experts and the statistical or empirical method determining educational mismatch as a level of education which is more than one standard deviation above or below the mean or modal value within a particular occupation. Each method has its advantages and weaknesses, see Sloane (2002) for a detailed discussion. However, different measures of educational mismatch are shown to be only weakly correlated, leading Sloane to conclude that measurement problems of educational mismatch are important. This highlights the importance of testing for the presence of a measurement error bias as discussed in Subsection 3.1. It is important to remember, however, that mismeasurement of the job-education-match is only a problem if the specialisation behaviour within households is correlated with it. This means, if the JEM of husbands is mismeasured in the same way for husbands being married to a very “traditional” women than for husbands with a wife sharing housework equally with them, the results are not biased.

Furthermore, it is worth noting that if the quality of the JEM at first appears to be a measure lacking precision when compared to a measures of labour market productivity such as the wage rate, it has also an advantage. For our research it is particularly interesting, that the JEM is unaffected by pure wage discrimination of employers, which is a problem so far unresolved in the research about the male marriage wage premia.

WHS: Labour market production is measured by weekly market work hours and the years of experience since the last long interruption of work (at least six months). Home production is measured by weekly hours of housework. In the survey, persons are asked how many hours they spend on average on housework (such as, washing, cooking or cleaning) on weekdays, on Saturdays, and on Sundays. These values are then summed up to the weekly amount of housework hours. These are less precise measures than time use survey data, which is only available for Switzerland for the years 1979/1980, as we will discuss in the next Section.

We report two measures of home production: first, the weekly housework hours without child care; and second, the total housework hours including child care. Spouses whose youngest child is between 14 and 18 years old and who give no information about child care are supposed to spend no time on it. Despite this assumption, the number of observations where no information about the hours of child care is given, is very high. The sample is reduced by about a third if the observations with missing information on child care are dropped. Therefore, the main sample is the one including information on housework hours only. But summary statistics and a subset of estimations are also presented for the reduced sample with information on total housework hours.

Sample definition: Within the total sample we have around 3000 individuals being married or living in a consensual partnerships, for which the relevant information about demographic variables, market and housework is available. Previously married persons are excluded from the sample as the within-household bargaining process is assumed to be more complex in these cases than the theoretical models suggest. Following our theoretical model, the men's sample is furthermore restricted to male employees, aged between 18 and 65 years old. For homogeneity reasons we include only Swiss citizens or foreigners with annual or residential permit. School leavers, domestic servants and employees in their own or in a relative's firm are excluded for similar reasons. Unfortunately, for many men and women there is no or only partial information available about their spouse. But for our analysis we can only use the observations of couples where information about market and housework hours as well as demographic information of both spouses is available. This leaves us with a sample of 1574 individuals living in 787 couples. A comparison of the main variables as shown in Table A1 (Appendix), however, does not hint at the presence of a substantial bias in the estimation sample due to this reduction in sample size.

4.2 Sample statistics

Table 1 reports the main characteristics of partners and households. The means and standard deviations are calculated taking account of sampling weights. In the total sample reported in the first column, 86% of all husbands of the sample have a good JEM, which is equivalent to a share of overeducated workers of 14%. This is at the lower boundary of the incidence of overeducation in international comparison [see Groot and Maassen van den Brink (2000)]. Firstly, educational

mismatch in Switzerland is low, at 15-20% for the total sample of employees depending on the definition of mismatch. Secondly, restricting the sample to married men or men living in a partnership reduces the average further [see Wirz and Atukeren (2004)].

< Table 1 >

Differences between spouses in main demographic and educational characteristics are rather small. Husbands are aged 42.2 years on average, wives only slightly younger at 39.8 years. The difference between the gender averages of education years is only one year. Despite this small difference in human capital endowments, spouses differ substantially in their respective market and home production. While married men work on average 44.1 hours a week in the labour market, women's average amounts to 16.8 hours. Similarly, market work experience of men is at 13.8 years on average, almost double the level of the women's level of 7.5 years. The average of women's housework hours is in contrast more than three times the level of men at 19.8 and 6.1 hours a week, respectively. Hence, these averages suggest that traditional specialisation of spouses, where the wife takes over the larger part of the housework and the husband concentrates on market work, is rather the norm in Switzerland. This results is confirmed by earlier and more recent research for Switzerland, see Buchmann, Kriesi, Pfeifer and Sacchi (2002), Widmer, Levy and Gauthier (2004) and Baumgartner and Fux (2004). However, there seems to be a considerable amount of heterogeneity between households. This becomes visible when looking at the high level of the standard deviations of the WHS variables (women's market hours and years of experience and the housework hours of both gender).

Comparison with other data sources: The same picture is shown by other data sources for Switzerland, like the Swiss Labour Force Survey (SLFS) of 2000, where a detailed questionnaire on "unpaid work" was added.¹⁰ But in the SLFS, the amount of housework is recorded to be substantially higher than in the SHP (1999), i.e. by about 10 hours on average for both genders. This represents 30% for women and even more than 100% for men. This might be due to the more restrictive definition of housework and the less detailed question used in the SHP than in the SLFS.¹¹ Strub and Bauer (2002) found comparable average hours to the SHP (1999) for weekly hours of housework, that is 7.6 hours for men and 21.6 hours for women, when only cleaning, cooking, washing and shopping is included. Hence, overall housework hours seem to be substantially underestimated by the SHP (1999) data. Bauer (1998, 2000) finds

¹⁰See Bundesamt für Statistik (1999, 2003).

¹¹In the special questionnaire of the SLFS of 1997 and 2000 the amount of time spent on 12 single activities of home production are asked one by one and then summed up. These activities include also taking care of garden and pets, driving kids somewhere and taking care of elderly persons. In the SHP persons are asked to give the total amount of hours they spend on housework, and only three activities are explicitly mentioned solely as examples (washing, cooking, cleaning).

similar results using a comparable question and methodology. Furthermore, he concludes that the correct amount of weekly housework hours of women is around 30-50% higher than the presented figure, when comparing earlier SLFS to the detailed time use survey data of 1979/1980 and 1994 in Switzerland and to similar survey data for Germany in 1991/1992. Strub and Bauer (2002) making similar comparisons find that even the SLFS data still underestimates the correct weekly housework hours of women by 3 hours. At the same time the men's housework hours in the SLF seem to overestimate the corresponding hours of men by also 3 hours.

Thus, we conclude from the above figures that the amount of housework done by women in the SHP (2000) data must be considered as being underestimated by at least 30%, due to the more restrictive definition of housework and the less precise methodology used in the survey. Due to the more restrictive definition of housework in the SHP some underestimation of men's total housework seems also reasonable to assume, although this underestimation should have quite a limited extent. This last conclusion is also supported when comparing total work hours for both genders. The total amount of work including home and market work as measured by the SHP (1999) is reported to be 37 hours for women and 50 hours for men. Such a large difference between women's and men's total working hours is not confirmed by the SLFS data for Switzerland, hence the underestimation of men's housework hours in the SHP data is probably not that high. These conclusions should be kept in mind in interpreting the results.

< Table 2 >

General household characteristics: The rather traditional preferences of Swiss couples are also highlighted by the fact that the clear majority of couples are married. Only about 9% of couples in the sample are living in consensual partnerships. Table 2. reports these households' characteristics. Furthermore, the sample is quite representative for Swiss households with respect to family composition. The majority of households (60%) have at home at least one child younger than 18 years old and in about 50% of these households the youngest child is aged between 0 and 6 years. Eight percent of couples have children older than 18 years living in their household and 12% of couples have children living outside their household. Finally, 20% of couples did not have children when the survey was taken. The second column documents the fact that in the sample restricted to couples where information on child care is not missing the couples without children are somewhat overrepresented, their share increasing from 20% to 36%. Similarly the share of consensual partnerships is higher. These facts will be taken care of in the estimation. For both measures of housework the difference between spouses average hours is substantial at around 12-13 hours a week. But again, the level of the standard deviation of these differences in both housework measures is quite high

Restricted sample: Restricting the sample to couples, where information on child care is available, does not alter the picture substantially. Overall, the

sample averages of age and labour market experience are slightly higher as the missing variable problem is by definition less important for older parents having older children. Linked to that, average hours of weekly housework are also somewhat lower, roughly two hours for women and half an hour for men. On average, including the hours for child care increases total housework hours by one hour for men and by one and a half hours for women.

These descriptive statistics show clearly that a high degree of specialisation of spouses between market and home production (WHS) is rather the norm in Switzerland. However, there seems to be also a substantial amount of variation in WHS between households, the impact of which on the husband's JEM quality will be investigated in the following analysis.

5 Descriptive evidence

Figures 1–3.4 and Tables 3.1 and 3.2 show the relationship between a husband's quality of his JEM and the degree of WHS measured by sample averages of spouses market and housework hours. Following the theory, we expect husbands with a “stay-at-home spouse” having a high degree of WHS to have a significantly better JEM quality than husbands whose spouses' are working full time in the market and are sharing housework more equally.

5.1 Using wives' market work as measure for WHS

We first measure the WHS using women's lower labour market participation as a proxy following the specification of equation (1a). For this purpose, we use two variables: 1) weekly market work hours and 2) years of work experience since the last interruption of at least six month. Market work hours emphasise the current degree of specialisation in housework versus market work. In contrast to that, experience rather measures the degree of labour market participation or specialisation over the past years of marriage and working life.

< Figures 1 and 2 >
< Table 3.1 >

Surprisingly and in contrast to the theory, market work hours of women are not negatively related with the frequency of a good JEM (“good match”) of their husbands in the total sample as shown in Figure 1. The only exception where a weak negative correlation is found is the case of households with children younger than 18 years old living in the household. When women's market work experience is taken as measure for WHS, as shown in Figure 2, the picture is similar. Only in the case of couples with children living in the household, the relationship between the wife's market work experience and the husband's JEM quality is slightly negative as expected.

5.2 Using spouses' housework hours as measure for WHS

< Figures 3.1 and 3.3 >

< Table 3.2 >

In a second step, following the specification of equations (1b) and (1c) WHS is measured directly by housework hours or the difference between spouses' housework hours (womens' minus mens' hours). In Figure 3.1 we see that the descriptive results are in line with theoretical expectations. Husbands with a good JEM are found (on average) to have wives who do more hours of housework than the wives of the husbands with a bad JEM. This is also the case when child care is included in household work, as shown in Figure 3.3.

< Figures 3.2 and 3.4 >

< Table 3.2 >

In Figure 3.2 and 3.4 both spouses' average housework hours are added in order to see which spouse's homework hours variation is largest, hence determine the variation of WHS. Clearly, this variation is mostly determined by the difference in average homework hours between the wives' of more successful husbands and the wives of less successful employees - independent of the presence of children in the household. Men's variation in housework hours contributes only to a minor extent to these differences between spouses' housework. It seems that the amount of hours spent in home production varies significantly only in the case of men living with children in the same household, where child care increases the hours spent on household work. This variation is negatively related to the quality of the JEM, as expected. But overall, the housework hours of men seem to be remarkably robust, around 7 hours a week. This is in line with results from studies on other comparable countries, e.g. Stratton (2003).

Most interestingly, spouses who do not have children younger than 18 years living in their household still report a substantial amount of hours spent on child care. This may be care given to children of neighbours, friends or family members, probably mostly grandchildren [see Bundesamt für Statistik (1999)]. This fact will be taken into account when interpreting the results.

Hence, a first look at the data shows no clear evidence for an effect of WHS on a husband's JEM when women's lower market work participation (weekly hours, experience) are taken as proxy for the degree of WHS. Nevertheless, ample evidence for this relationship is found when taking instead the respective housework hours of spouses as a direct measure. Moreover, based on these simple sample averages, it seems that mainly the women's housework hours contribute to the positive correlation of the difference between spouses' housework hours and the quality of the husband's JEM. These initial findings from descriptive data analysis speak against the model of Becker (1985) and lend support rather for the theory of within-household interaction of François (1998).

6 Estimation results

In this section we test for the impact of WHS on the quality of the JEM of husbands using probit estimation following the empirical approach outlined in Section 3. Sampling weights adjusted for design and non-response bias are taken into account in the estimation. These estimations will be done for the different measures of WHS discussed in the previous Sections.

6.1 Using wives' market work as measure for WHS

In Table 4 the results of the probit model using women's weekly market working hours and experience as measure for WHS (equation (1a)) are shown.

< Table 4 >

Surprisingly the working hours of women (column A) have no impact on a husband's probability to have a good JEM in the total sample, confirming the result from the initial descriptive analysis. This stands in contradiction to previous research on specialisation effects (or marriage wage premia) where the impact of women's working hours on a husband's wage is a confirmed empirical fact. More in line with previous research, the level of women's working experience (column B) has a significant negative impact. We tested whether the experience variable simply captured a cohort effect by adding women's age directly to the probit equation. The results were unaltered. Therefore, as expected, a husband's probability of having a good JEM is higher if his spouse has an interrupted working career and thus less years of working experience. A coefficient of -0.002 implies that the difference in this probability is around 4%, when comparing couples of otherwise identical characteristics, for instance aged around forty years, but where the wife in one case interrupted her working career, e.g. when getting married, and in the other case the wife has 20 years of uninterrupted working experience.

The Rivers-Vuong test indicates by the p -values (based on the z -statistic of the added residual) that the null hypothesis of exogeneity of the WHS variables ($H_0 : \rho_1 = 0$) cannot be rejected in the case of working hours or in the case of experience taken as a measure for WHS at conventional levels of significance. The significance level shows the probability that a given result is only due to chance when we think it is true, and such high levels are of course unacceptable. This test was done in a set-up where women's market work hours and experience are explained in a first step as a function of her age and education, the number and age of children, her parents' education and social status and the women's attitude towards public measures favouring equal chances for both genders and the importance of national traditions. In a second step, equation (3a) is estimated taking into account the residual of the first step, see Section 3.1 for a more detailed description of this test. The fact that education is measured by dummies (nine classes) increases the number of instruments artificially. However, the guideline suggested by Staiger and Stock (1997) to have 10-20 observations

per instrument is still respected. In order to evaluate the quality of the instruments we perform a joint test of the significance of the instrumental variables in the first-stage regression. The F -statistic of this test is clearly satisfactory for the total sample. Hence, we conclude that sorting of more able husbands with “traditional” spouses or measurement error of the quality of the JEM of the husband do not bias the results of the one-stage probit estimation. Therefore, the *Model 1* applies when testing the impact of labour market participation variables of wives as measure for specialisation

In a further step, we investigate whether these results are sensitive to the amount of housework to be shared between spouses. Therefore, estimations are run separately for couples living in households with or without dependent children (younger than 18 years), respectively. It seems reasonable to assume *a priori* that the amount of home production that cannot be substituted by market products and services (or not in a satisfactory way) differs substantially between the two samples. Particularly, housework linked to child raising and child care is itself often very time consuming and parents may frequently have a preference for home production even if market substitutes are available.

The figures in column (C)-(F) show that the results are almost the same in two subsamples of households. Due to the smaller size of the subsamples, the efficiency of the estimate is lowered as shown by the lower z -statistics of the coefficients and the F -statistic of the instruments’ test. Hence, the impact of WHS is observed to be quite robust to variations in the total amount of housework, which means that the results can be generalised to any kind of household independent of its composition.

Overall, these results show evidence for WHS in explaining a substantial part of husbands’ JEM quality, when women’s working experience is taken as proxy for WHS. Taking women’s weekly working hours as proxy does not confirm this relationship. However, this fact can also be attributed to the high variation of women’s working time observed over their working life on average in Switzerland, making this measure of specialisation very unprecise.

6.2 Using spouses’ housework hours as measure for WHS

In a second step, we test for the impact of WHS on the JEM of the husband measured directly by both spouses’ home production. As discussed before, this allows us first to measure WHS more precisely and to discriminate between different models of within-household interaction.

< Table 5 >

In Table 5, column (A) the results of the probit estimation including both spouses housework hours as individual variables as formulated in equation (1b) are presented. Most interestingly, women’s housework hours are shown to have a significant positive and independent impact on a husband’s probability to have a good JEM. Husbands’ housework hours, on the contrary, do not seem to have a significant impact on their own JEM, albeit the exact significance

level of the coefficient's z -statistic is close to 10%. Overall this results confirms the expectation that WHS improves the JEM of a husband. The specialisation of a wife in home production seems to improve the chances of her husband to have a good job. Furthermore, this result clearly speaks against the hypothesis of Becker's model suggesting that WHS has its effect exclusively through the lowering of the husband's housework burden. This result is, however, in line with the results of Hersch and Stratton (2000).

Taking the difference between spouses' housework hours as measure for WHS (column (B)) confirms these conclusions. This is the specification formulated in equation (1c). A significant positive impact of this difference on the JEM of husband's is found. The statistical significance of the results is clearly improved when compared to the previous specifications (1a, 1b). This confirms the importance of taking into account both spouses housework hours for approximating WHS. The impact of specialisation on husbands' productivity seems clearly related to the degree the housework burden is shared between spouses and not to the number of housework hours a husband does not do anymore when getting married.

The p -value of the Rivers-Vuong test being above the 50% level again indicates clearly that the null hypothesis of exogeneity of the specialisation variables ($H_0 : \rho_1 = 0$) cannot be rejected. The difference between housework hours of spouses is a continuous variable, which is not censored on either side like working hours and experience of spouses. Hence, it is estimated in the first step of the Rivers-Vuong test procedure with ordinary least squares. Thus we have no evidence that the results are biased by a sorting mechanism of more able husbands and "traditional" women or measurement error.

The main results are also shown to be robust to variations in the amount of housework accruing to a household as summarized in columns (C)-(F) for the subsamples with and without dependent children living in the household, although the efficiency of the estimates is again lowered due to the smaller sample size. As an exception to this, the z -statistic of the coefficient of women's housework hours for families with dependent children dropped just below the significance level of 10%, but its size is unaltered. Intuitively housework without hours spent on child care is not the most appropriate measure for WHS in this subsample. This explanation will be confirmed subsequently. Partly due to the smaller sample sizes the F -statistic of the instruments' test is rather at the lower boundaries. In general, the impact of WHS is quite robust to variations in the amount of housework accruing to the household overall.

These estimations are repeated by taking total housework hours, including child care for measuring WHS. The results are presented in Table 6.

< Table 6 >

The picture is almost identical when the results for the total sample and for household with dependent children are looked at (columns (A)-(D)). The statistical significance of the estimated coefficients is even improved. Nevertheless, contrary to the previous results, the impact of housework of women in families

with dependent children is highly significant and at the upper boundaries of the range of estimates. This confirms the strong and independent impact of women’s housework. The results for couples without dependent children living in their household are less clear-cut in this case (columns (E)-(F)). Including child care, which is probably given to children of neighbours, friends or family members, does not seem to have the same impact as housework or child care given to own children. This result could be interpreted in the sense that such housework is considered a rather optional activity by the spouses and is not subject to a bargaining process within households.

Additional robustness checks were done adding dummies for women’s education as proxy for her potential earnings directly to the probit estimation following the argumentation of Skatun (2004). The results are shown in Table A2 in the Appendix. Again the size and significance of the impact of WHS on the JEM of the husbands are unaltered by these additional controls. Hence, the impact of WHS measured by spouses’ housework hours is not biased by a possible impact of potential earnings of the wife on her husband’s bargaining power at work.

Simulation of the quantitative impact:

< Figure 4 >

< Table 7 >

Figure 4 presents the quantitative implications of the results found. More precisely, we see in this figure how the predicted probability of a husband of having a good JEM varies with the degree of WHS when approximated by the difference between spouses housework hours. In order to estimate this impact, we estimate the probability for a husband, with average values in explanatory variables (see Table 7) and differing degrees of WHS. In the first case, the wife is not specialising in housework at all, hence the difference between their housework hours equals zero. Then, this difference is increased to 15 and 30 hours respectively. This range represents roughly the variation of one standard deviation (12.4 hours) above and below the mean of this difference in weekly work hours (13.7) hours, observed in our sample. In that context, it is important to remember that the correct amount of housework hours in the SHP (2000) is underestimated by 30-50% for women. Thus, we conclude that women indicating in the survey to do 30 hours of household work more than their husbands may actually work 40 hour more, being fully specialised in home production. The predicted probability of having a good JEM increases from 89% to 94% overall following the two increases of 15 hours each ranging from no specialisation at all to complete specialisation of the spouse in housework.

Hence, when WHS is measured by the amount of housework production the wife does in excess of a husband’s own housework, it is clearly shown to have a noticeable positive impact on a his probability to have a good JEM. Given the average level of the difference between spouses’ housework of 15 hours, the average impact of WHS should be around 3 percentage points. The size of this

impact is estimated under the assumption of a standard normal distribution of the error term in our model. 3 percentage points represent around 30% of the size of the marginal effect of marriage on men’s probability to have a good JEM, estimated at 9.5 percentage points by Wirz and Atukeren (2004). Hence, a substantial share of the impact of marital status on the JEM of men can indeed be attributed to WHS effects.

However, since these results and tests are based on cross-section estimates, it is not possible to infer directly the existence of causal relations between the independent variables and the dependent variable. Nevertheless, the regularities described are robust and could possibly indicate the direction of causal links. At least, empirical evidence does not seem to contradict our hypotheses. However, further research based on data from a panel covering a sufficiently long time period is needed to confirm these results.

7 Conclusion

We indeed find support for the common idea cited in the introduction, that a “great woman” may help a man to be “great”. A “great woman” is defined in this context admittedly simplifying as being a woman taking over a larger share of the housework burden. Overall, the results show clearly that such *within-household specialisation* (WHS) explains a substantial part of the improved *job-education-match* (JEM) of husbands, whether a wife’s labour market participation (experience) or the difference between spouses’ housework hours are used to measure WHS. Testing for a possible endogeneity bias’ we do not find evidence for a sorting process of more able husbands with “traditional” spouses to bias this result. However, further research based on panel data of spouses market and home production, covering a sufficiently long time span, is needed to test whether these results based on cross-sectional data really show a true causal relationship.

Nevertheless, the empirical evidence so far strongly hints at an independent and significant impact of women’s housework hours on her husband’s JEM. The impact of WHS seems clearly related to the amount of housework hours taken over by the wife independent of the husband’s number of housework hours. This clearly speaks against the explanation of Becker (1985) but for the theory of François (1998) stipulating that the work motivation of a husband is increased by his gains from within-household trade. These conclusions are robust to various assumptions about the bargaining process between spouses.

These findings imply that the impact of education policy is affected by social and economic factors determining the time allocation decisions of spouses. More precisely, measures favouring a traditional specialisation of spouses between market and home production help improve the allocation of husbands to jobs matching their education. However, it is important to note in this context that this improvement in allocation is acquired at the cost of more frequently interrupted working careers of women on average, hence their lower involvement in the labour market. Thus, if household specialisation indeed improves the

JEM of husbands, then social policies favouring “traditional” marriages clearly favour the utilisation of men’s human capital to the detriment of the utilisation of women’s capital.

8 Annexes

Tables 1-7.

Figures 1-4.

8.1 Appendix

Table A1-A2.

9 References

References

- [1] Bauer, T. K. (2002); “Educational mismatch and wages: a panel analysis”, *Economics of Education Review*, 21(3), pp. 221-29.
- [2] Bauer, T. (2000); *Die Familienfalle*, Rüegger, Chur / Zürich.
- [3] Bauer, T. (1998); “The impact of family structure on time use and potential wage in Switzerland”, *International Journal of Manpower*, 19 (7), pp. 507-519.
- [4] Baumgartner, D.; Fux, B. (2004); “Und sie bewegen sich doch nicht: die Männer. Zur geschlechtsspezifischen Verteilung der Erwerbsarbeit in Familien”, in Zimmermann, E. and Tillmann, R., “Vivre en Suisse 1999-2000”, Peter Lang, Bern, pp. 109-130.
- [5] Becker, G. S. (1985); “Human capital, effort and the sexual division of labor”, *Journal of Labour Economics*, 3 (1), pp. S33-S58.
- [6] Becker, G. S. (1991); “A Treatise on the Family”, Harvard University Press, Cambridge, MA.
- [7] Bonke, J.; Datta Gupta, N.; Smith, N. (2003); “Timing and flexibility of housework and men and women’s wages”, IZA Discussion Paper No. 860, Institute for the Study of Labor (IZA), Bonn.
- [8] Büchel, F.; Battu, H. (2002); “The theory of differential overqualification - Does it work?”, IZA Discussion Paper No. 511, Institute for the Study of Labor (IZA), Bonn.
- [9] Büchel, F.; van Ham, M. (2003); Buchmann, M.; Kriesi, I.; Pfeifer, A.; Sacchi, S. (2002); Overeducation, regional labour markets and spatial flexibility”, *Journal of Urban Economics*, 53 (3), pp. 482-493.

- [10] Buchmann, M.; Kriesi, I.; Pfeifer, A.; Sacchi, S. (2002); "Halb drinnen - halb draussen. Analysen zur Arbeitsmarktintegration von Frauen in der Schweiz", Rüegger, Chur / Zürich.
- [11] Chevalier, A. (2003); "Measuring over-education", *Economica*, 70 (279), p. 509-531.
- [12] Bundesamt für Statistik (1999); "Unbezahlt - aber trotzdem Arbeit", BFS, Neuchâtel.
- [13] Bundesamt für Statistik (2003); "Auf dem Weg zur Gleichstellung? Frauen und Männer in der Schweiz: dritter statistischer Bericht", BFS, Neuchâtel.
- [14] Dolton, P.; Silles, M. (2001); "Over-education in the graduate labour market: some evidence from alumni data", CEE Discussion paper 9, June 2001.
- [15] François, P. (1998); "Gender discrimination without gender difference: theory and policy responses", *Journal of Public Economics* 68 (1), pp. 1-32.
- [16] Frank, R. H. (1978); "Why women earn less: the theory and estimation of differential overqualification", *American Economic Review*, 68(3), pp. 360-373.
- [17] Hartog, J. (2000); "Over-education and earnings: where are we, where should we go?", *Economics of Education Review*, 19 (2), pp. 131-147.
- [18] Hersch, J.; Stratton, L. S. (2000); "Household specialisation and the male wage premium", *Industrial and Labor Relations Review*, 54 (1), pp. 78-94.
- [19] Manser, M.; Brown, M. (1980); "Marriage and household decision-making: a bargaining analysis", *International Economic Review*, 21 (1), pp. 31-44.
- [20] McElroy, M. B.; Horney, M. J. (1981); "Nash-bargained household decisions: toward a generalization of the theory of demand", *International Economic Review*, 22 (2), pp. 333-349.
- [21] McElroy, M. B.; Horney, M. J. (1990); "Nash-bargained household decisions: reply", *International Economic Review*, 31 (1), pp. 237-42.
- [22] McGoldrick, K. M.; Robst, J. (1996); "Gender differences in overeducation: a test of the theory of differential overqualification", *American Economic Review*, 86 (2), pp. 280-84.
- [23] Green, F.; MacIntosh, S. (2002); "Is there a genuine underutilisation of skills amongst the overqualified?", Paper presented at the International Conference on Over-education in Europe: What Do We Know?, 22-23 November 2002, Berlin.
- [24] Sousa-Poza, A., Schmid, H.; Widmer, R.; "The allocation and value of time assigned to housework and child-care: an analysis for Switzerland", *Journal of Population Economics*, 14, pp. 599-618.

- [25] Sloane, P. J. (2002); "Much ado about nothing? What does the over-education literature really tell us?", Keynote Address, International Conference on Over-education in Europe: What Do We Know?, 22-23 November 2002, Berlin.
- [26] Sloane, P. J.; Battu, H.; Seaman, P. T. (1996); "Overeducation and the formal education / experience and training trade-off", *Applied Economics Letters*, 3 (8), pp. 511-15.
- [27] Skatun, J. D. (2004); "Behind every well paid married man: the impact of the partner's earning opportunity", *Australian Economic Papers*, 43 (1), pp. 1-9.
- [28] Staiger, D.; Stock, J. H. (1997); "Instrumental variables regression with weak instruments", *Econometrica*, 65 (3), pp. 557-586.
- [29] Stratton, L. S. (2003); "The division of work in the household", in Moe, K. S., "Women, Family and Work. Writings on the Economics of Gender", Blackwell, Oxford, Melbourne, Berlin, pp. 65-83.
- [30] Strub, S.; Bauer, T. (2002); "Wie ist die Arbeit zwischen den Geschlechtern verteilt?", Eidgenössisches Büro für die Gleichstellung von Mann und Frau, Bern.
- [31] Van der Meer, P. H., Batenburg, R. (2002); "Overeducation in the Netherlands: new trends old problems?", Paper presented at the International Conference on Over-education in Europe: What Do We Know?, 22-23 November 2002, Berlin.
- [32] Widmer, E.; Levy, R.; Gauthier, J.-A. (2004); "L'Implication dans les champs domestique et professionnel selon les phases de la vie familiale", in Zimmermann, E. and Tillmann, R., "Vivre en Suisse 1999-2000", Peter Lang, Bern, pp. 95-108.
- [33] Wirz, A.; Atukeren, E. (2004); "Overeducation in the Swiss labour market - Does anything go wrong?", Arbeitspapier/Working paper No. 85, KOF/ETH, Zurich.
- [34] Wooldridge, J. M. (2002); "Econometric Analysis of Cross Section and Panel Data", The MIT Press, Cambridge, MA.

Table 1. Descriptive statistics of spouses, mean (standard deviation)

	Total sample	Restricted sample: information on childcare available
<i>Husband's characteristics:</i>		
Good job-education-match (GJEM) (1)	0.865	0.878
Age	42.15 (10.34)	45.17 (11.51)
Experience (2)	20.81 (11.92)	23.91 (13.16)
Market work hours (HMWH), weekly	44.14 (9.03)	43.56 (9.02)
Education years	13.80 (2.46)	13.74 (2.34)
Tenure	11.21 (10.32)	12.85 (11.54)
Non-Swiss nationality	0.182	0.136
Housework hours (HHWH), weekly	6.10 (5.72)	5.77 (5.37)
Total housework hours, including child care (THHWH), weekly	-	6.95 (7.44)
<i>Wife's characteristics:</i>		
Age	39.80 (10.20)	42.85 (11.54)
Experience (2)	7.53 (9.51)	9.86 (10.64)
Market work hours (WMWH), weekly	16.84 (16.38)	21.41 (16.92)
Education years	12.79 (2.37)	12.75 (2.28)
Non-Swiss nationality	0.167	0.104
Housework hours (WHWH), weekly	19.77 (10.91)	17.61 (10.55)
Total housework hours, including child care (TWHWH), weekly	-	19.17 (11.84)
Number of observations (couples) (3)	787	439

Source: Swiss Household Panel (SHP) 1999

Notes:

- (1) GJEM is a dichotomous variable taking value 1 when a husband's educational qualifications correspond to his job and 0 if his qualifications are superior to his job. This notation is inverse to the notation used in the overeducation literature. Generally the overeducation variable takes the value 1 when a person is working overeducated and 0 otherwise.
- (2) Years since last long interruption of market work (min. 6 month).
- (3) The sample includes married or in consensual partnerships living male employees aged between 18 and 65 years, who are Swiss or foreigners with an annual or a permanent residential permit, and have at least obtained compulsory schooling. The reference persons are Swiss, have obtained vocational training as highest education degree and do not have children.

Table 2. Descriptive statistics of households, mean (standard deviation)

	Total sample	Restricted sample: information on childcare available
<i>Household characteristics:</i>		
Consensual partnership	0.087	0.145
Child / children living in household (HH) younger than 18 years old	-	-
Child / children living in household (HH) between 0 and 6 years old	0.593	0.281
Child / children living in household (HH) between 7 and 13 years old	-	-
Child / children living in household (HH) between 14 and 17 years old	0.336	0.106
Child / children living in household (HH) older than 18 years old	-	-
Child / children living outside household (HH)	0.180	0.039
No children	-	-
	0.077	0.136
	-	-
	0.080	0.143
	-	-
	0.124	0.219
	-	-
	0.203	0.357
	-	-
Difference between spouses' housework hours, weekly (WHWH-HHWH)	13.67 (12.43)	11.84 (12.10)
Difference between spouses's total housework hours, including child care, weekly, (TWHWH-THHWH)	-	12.22 (13.64)
Number of observations (couples) (1)	787	439

Source: Swiss Household Panel (SHP) 1999

Notes:

(1) Sample definition see Table 1.

Figure 1. Average market work hours of women by quality of the job-education-match of their husband

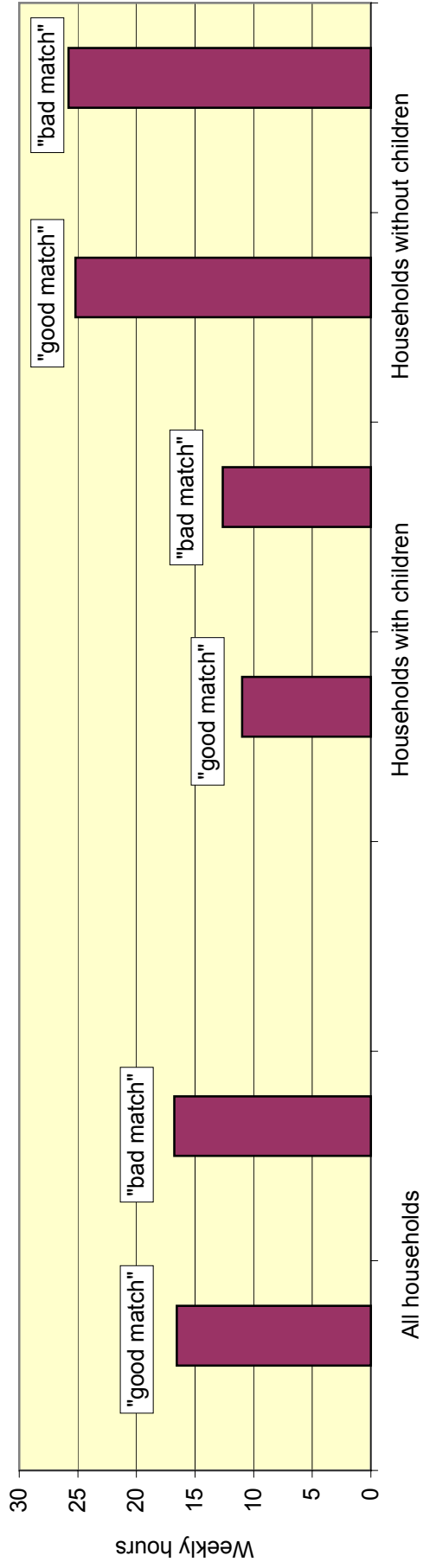


Figure 2. Average market work experience of women by quality of the job-education-match of their husband

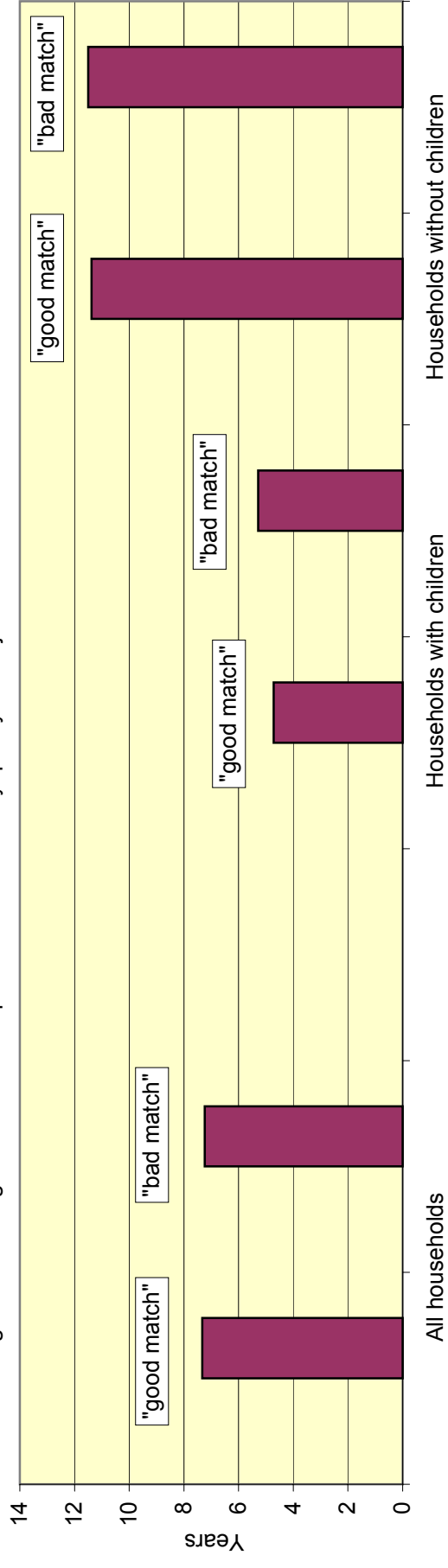


Figure 3.1 Average difference between spouses' housework hours by quality of the job-education-match of the husband

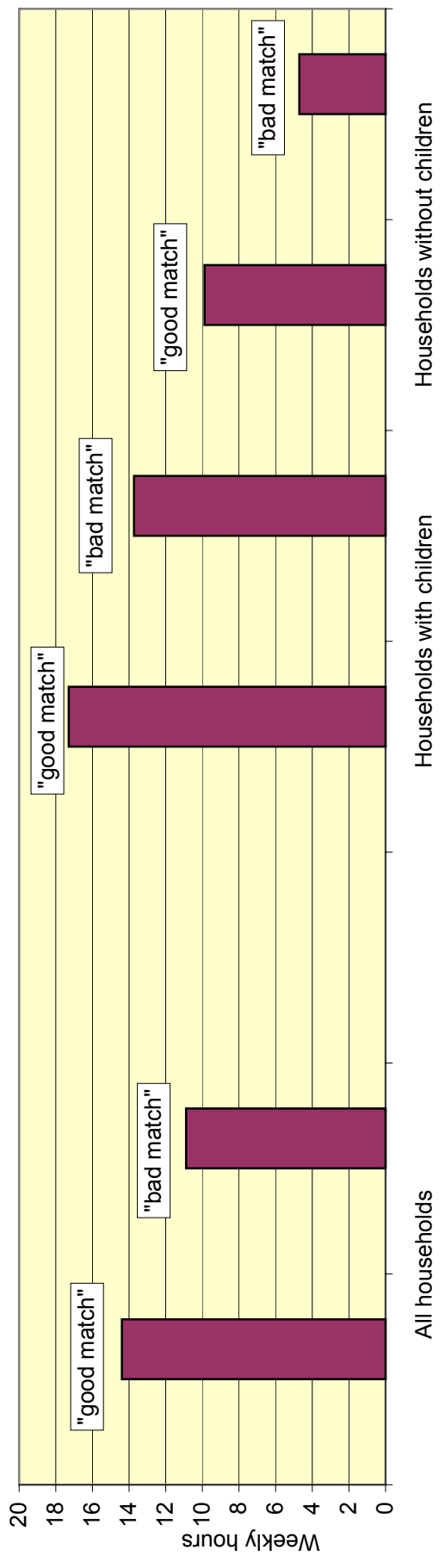


Figure 3.2 Average difference between spouses' housework hours by quality of the job-education-match of the husband

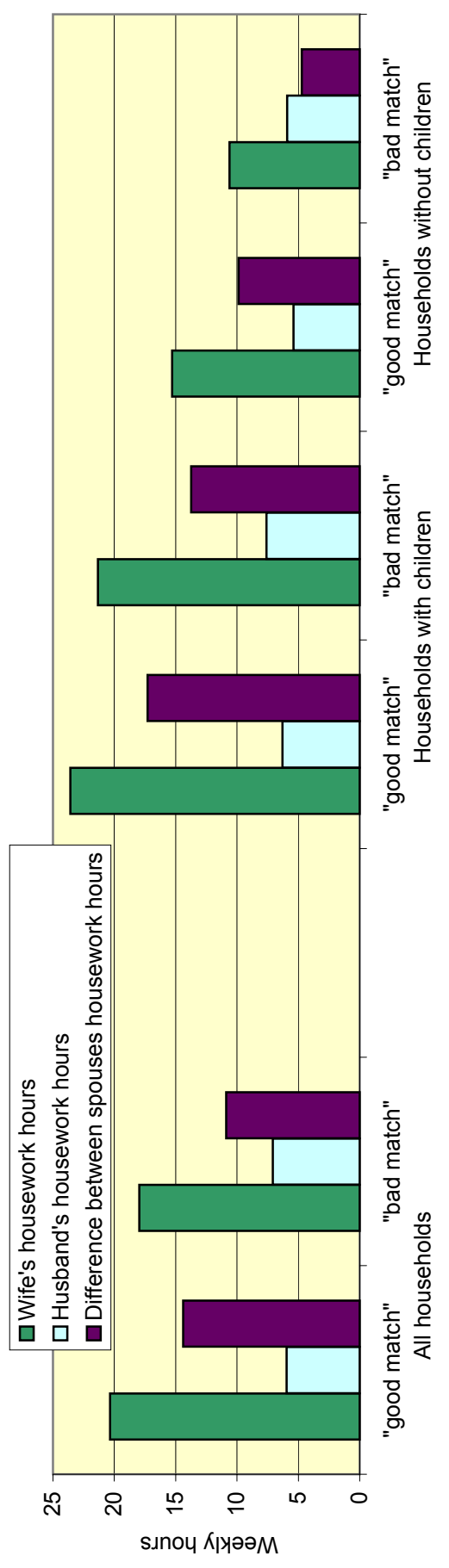


Figure 3. 3. Average difference between spouses' total housework hours by quality of the job-education-match of the husband (incl. child care hours, restricted sample)

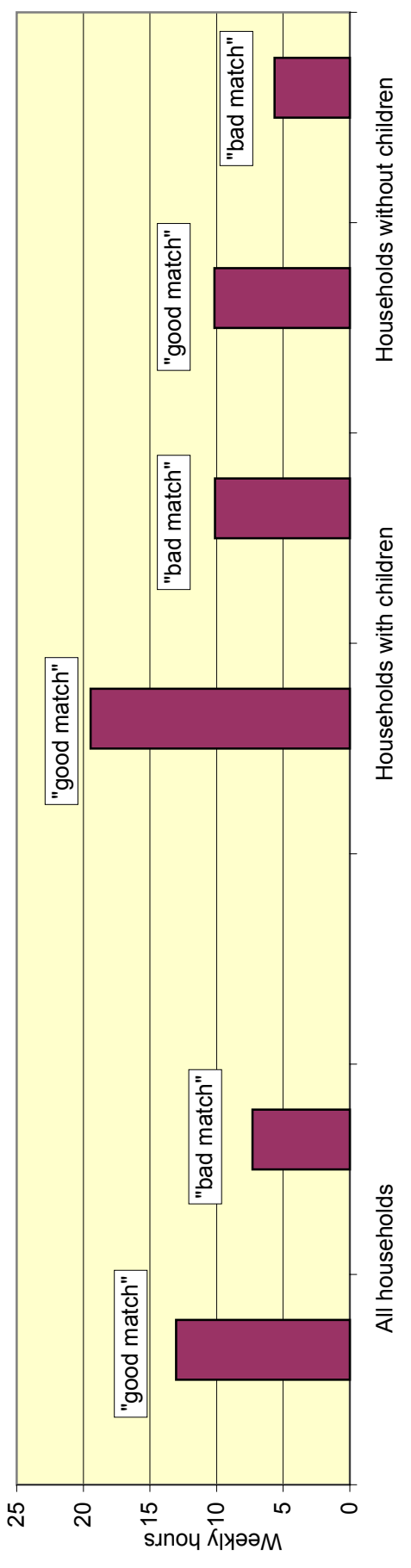


Figure 3.4 Average difference between spouses' total housework hours by quality of the job-education-match of the husband (incl. child care hours, restricted sample)

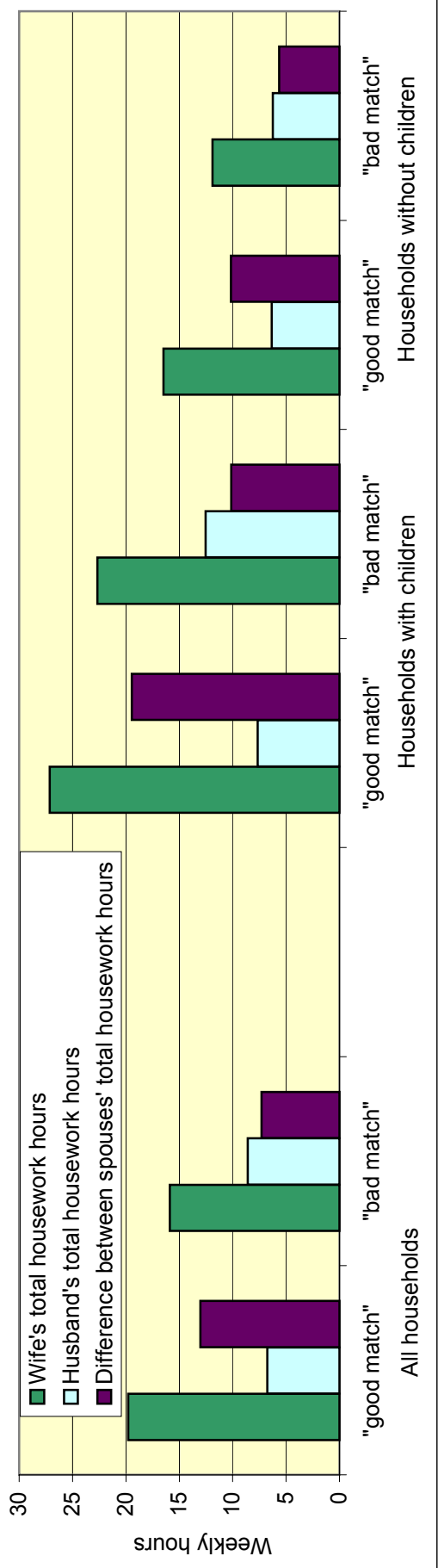


Table 3.1 Within-household specialisation (WHS) and husband's job-education-match (JEM) - total sample

Good job-education-match (GJEM)	All households		Households with dependent children (4)		Households without dependent children (5)	
	Yes (2)	No (3)	Yes (2)	No (3)	Yes (2)	No (3)
Mean of WHS measures: (1)						
Wife's market work hours (weekly)	16.56 (16.32)	16.78 (16.29)	10.99 (13.35)	12.63 (14.06)	25.21 (16.75)	25.79 (17.34)
Wife's market work experience (years)	7.33 (9.23)	7.24 (9.28)	4.72 (6.94)	5.28 (7.23)	11.38 (10.77)	11.5 (11.68)
Wife's housework hours (WHWH), weekly	20.34 (11.21)	17.96 (10.01)	23.59 (11.09)	21.34 (9.17)	15.28 (9.38)	10.62 (7.60)
Husband's housework hours (HHWH), weekly	5.95 (5.74)	7.07 (6.35)	6.30 (6.09)	7.61 (6.96)	5.40 (5.12)	5.91 (4.63)
Difference between spouses' housework hours (WHWH-HHWH)	14.39 (12.77)	10.89 (11.56)	17.29 (12.76)	13.73 (11.73)	9.88 (11.41)	4.71 (8.42)
Number of observations	679	108	413	74	266	34
% of total	86.28	13.72	52.48	9.40	33.80	4.32

Source: Swiss Household Panel (SHP) 1999

Notes:

- (1) Standard deviations in parentheses.
- (2) GJEM=1, educational qualifications correspond to worker's job.
- (3) GJEM=0, educational qualifications are superior to worker's job (overeducated worker).
- (4) Families with children younger than 18 years old, living in the same household.
- (5) Couples without children, with children living outside their household or with children older than 18 years old, living in the same household.

Table 3.2. Within-household specialisation (WHS) and husbands' job-education-match (JEM) - sample restricted to couples with information on child care

	All households		Households with children (4)		Households without children (5)	
	Yes (2)	No (3)	Yes (2)	No (3)	Yes (2)	No (3)
Good job-education-match (GJEM)						
Mean of WHS measure (1)						
Wife's total housework hours (TWHWH), weekly (6)	19.78 (12.15)	15.89 (10.60)	27.12 (11.76)	22.67 (10.15)	16.50 (10.83)	11.90 (8.77)
Husband's total housework hours (THHWH), weekly (6)	6.75 (6.98)	8.58 (10.28)	7.67 (7.96)	12.54 (14.89)	6.34 (6.46)	6.25 (5.20)
Difference between spouses' total housework hours (TWHWH-THHWH) (6)	13.03 (13.86)	7.31 (13.15)	19.45 (14.07)	10.13 (18.47)	10.16 (12.78)	5.65 (8.58)
Number of observations	385	54	119	20	266	34
% of total	87.70	12.30	27.11	4.56	60.59	7.74

Source: Swiss Household Panel (SHP) 1999

Notes:

(1) - (5) see Table 3.1.

(6) Total housework hours, including child care.

Table 4. The impact of within-household specialisation (WHS) - measured by wives' market work hours and experience - on husbands' job-education-match (JEM)

Probit estimates, marginal effects are presented	All households					
	(A)	(B)	(C)	(D)	(E)	(F)
<i>Dependent variable: Good job-education-match (GJEM) (1) (2)</i>						
Wife's market work hours (weekly)	-0.000 (0.14)	-	-0.002 (1.51)	-	0.001 (0.64)	-
Wife's market work experience (years)	-	-0.002 (1.88)*	-	-0.004 (1.67)*	-	-0.003 (1.63)
Number of observations (3)	787	787	457	457	300	300
Log-likelihood value	-283.6	-281.9	-182.6	-182.5	-94.4	-93.1
Pseudo R-squared	0.093	0.099	0.099	0.099	0.125	0.136
<i>Rivers-Vuong test for exogeneity:</i>						
p-value (4)	0.607	0.814	0.554	0.931	0.693	0.398
<i>F-Test of validity of instruments (5)</i>						
(p, n-k)	6.94 (34, 614)	5.27 (34, 614)	2.84 (20, 376)	2.89 (20, 376)	3.12 (22, 216)	3.14 (22, 216)
p-value	0.000	0.000	0.000	0.000	0.000	0.000

Robust z-statistics in parentheses, * significant at 10%

Source: Swiss Household Panel (SHP) 1999

Notes:

- (1) GJEM is a dichotomous variable taking value 1 when a husband's educational qualifications correspond to his job and 0 if his qualifications are superior to his job.
- (2) Also included in the regression are the husband's education, age, age squared, market work experience, tenure, health, dummies for part-time working, for a Non-Swiss nationality, for difficulties in professional or private life and for persons being not married but living in a consensual partnership. Sample definition see Table 1.
- (3) 30 observations are lost in the estimation of the model in the subsample of households with dependent children (column (C)-(D)) due to the lack of variation of the dependent variable within the education categories "compulsory schooling" and "general training".
- (4) p-value of estimated residual of the reduced form for working hours and experience respectively, see Section 3.1. A high p-value means the null-hypothesis of exogeneity cannot be rejected.
- (5) The term instrument is used for the explanatory variables of the WHS measure in the first step of the River-Vuong test. A low p-value means the null-hypothesis that none of these variables has an effect on the explained WHS measure has to be rejected. The variables used are age, education and nationality of the wife, her attitude towards tradition, her non-wage income, the number of children, six dummies for the age of children living inside or outside the household, nationality of mother, education of parents and social status of fathers' jobs (Treiman's prestige scale). The exogeneity of these variables was tested adding them to the probit of the husband's JEM and performing a simple significance test (based on the z-statistics).

Table 5. The impact of within-household specialisation (WHS) – measured by spouses' housework hours - on husbands' job-education-match (JEM)

Dependent variable: Good job-education-match (GJEM) (1) (2)	All households		Households with dependent children		Households without dependent children	
	(A)	(B)	(C)	(D)	(E)	(F)
Wife's housework hours (WHW)	0.002 (1.65)*	-	0.002 (1.56)	-	0.004 (1.74)*	-
Husband's housework hours (HHW)	-0.002 (1.40)	-	-0.003 (1.13)	-	-0.002 (0.66)	-
Difference in housework hours (DHW= WWH-HHW)	-	0.002 (2.13)**	-	0.003 (1.93)*	-	0.003 (2.00)**
Number of observations (3)	787	787	457	457	300	300
Log-likelihood value	-281.4	-281.5	-182.1	-182.1	-92.5	-92.6
Pseudo R-squared	0.100	0.100	0.101	0.101	0.142	0.141
<i>Rivers-Vuong test for exogeneity:</i>						
<i>p</i> -value (4)	-	0.507	-	0.580	-	0.659
<i>F</i> -Test of validity of instruments (5)						
(<i>p</i> , <i>n</i> - <i>k</i>)	-	6.14 (34, 651)	-	2.49 (31, 389)	-	3.87 (33, 217)
<i>p</i> -value	-	0.000	-	0.000	-	0.000

Robust z-statistics in parentheses, * significant at 10%, ** significant at 5%

Source: Swiss Household Panel (SHP) 1999

Notes:

- (1) GJEM is a dichotomous variable taking value 1 when a husband's educational qualifications correspond to his job and 0 if his qualifications are superior to his job.
- (2) Also included in the regression are the husband's education, age, age squared, market work experience, tenure, health, dummies for part-time working, for a Non-Swiss nationality, for difficulties in professional or private life and for persons being not married but living in a consensual partnership. Sample definition see Table 1.
- (3) 30 observations are lost in the estimation of the model in the subsample of households with dependent children (column (C)-(D)) due to the lack of variation of the dependent variable within the education categories "compulsory schooling" and "general training".
- (4) *p*-value of estimated residual of the reduced form for working hours and respectively experience, see Section 3.1. A high *p*-value means the null-hypothesis of exogeneity can not be rejected.
- (5) The term instrument is used for the explanatory variables of the WHS measure in the first step of the River-Vuong test. A low *p*-value means the null-hypothesis that none of these variables has an effect on the explained WHS measure has to be rejected. The variables used are age, education and nationality of the wife, her attitude towards tradition, her non-wage income, the number of children, six dummies for the age of children living inside or outside the household, nationality of the mother, education of parents and social status of fathers' jobs (Treiman's prestige scale). The exogeneity of these variables was tested adding them to the probit of the husband's JEM and performing a simple significance test (based on the z-statistics). The optimal choice of instruments was adapted to the subsamples of households.

Table 6. The impact of within-household specialisation (WHS) – measured by spouses' total housework hours - on husbands' job-education-match (JEM)

Probit estimates, marginal effects are presented	All households	Households with dependent children	Households without dependent children
<i>Dependent variable:</i> Good job-education-match (GJEM) (1) (2)	(A)	(C)	(E)
	(A)	(C)	(E)
Wife's total housework hours (TWHW) (3)	0.002 (1.76)*	0.005 (2.49)**	0.003 (1.48)
Husband's total housework hours (3) (THHW)	-0.002 (1.31)	-0.002 (1.21)	-0.001 (0.31)
Difference in total housework hours (3) (DTHW= TWWH-THHW)	-	-	-
	0.002 (2.23)**	0.003 (2.92)***	0.002 (1.60)
Number of observations (4)	439	131	300
Log-likelihood value	-144.0	-39.3	-93.4
Pseudo R-squared	0.117	0.281	0.134
<i>Rivers-Vuong test for exogeneity:</i>			
<i>p</i> -value (7)	-	-	-
<i>F</i> -Test of validity of instruments (5)			
(<i>p</i> , <i>n</i> - <i>k</i>)	-	-	-
<i>p</i> -value	4.01 (34, 337)	8.22 (30, 76)	3.19 (33, 217)
	0.000	0.000	0.000

Robust z-statistics in parentheses, * significant at 10%, ** significant at 5%, *** significant at 1%

Source: Swiss Household Panel (SHP) 1999

Notes:

- (1) GJEM is a dichotomous variable taking value 1 when a husband's educational qualifications correspond to his job and 0 if his qualifications are superior to his job.
- (2) Also included in the regression are the husband's education, age, age squared, market work experience, tenure, health, dummies for part-time working, for a Non-Swiss nationality, for difficulties in professional or private life and for persons being not married but living in a consensual partnership. Sample definition see Table 1.
- (3) Total housework hours, including child care.
- (4) 8 observations are lost in the estimation of the model in the subsample of households with dependent children (column (C)-(D)) due to the lack of variation of the dependent variable within the education categories "compulsory schooling" and "general training".
- (5) *p*-value of estimated residual of the reduced form for working hours and experience respectively, see Section 3.1. A high *p*-value means the null-hypothesis of exogeneity can not be rejected.
- (6) The term instrument is used for the explanatory variables of the WHS measure in the first step of the River-Vuong test. A low *p*-value means the null-hypothesis that none of these variables has an effect on the explained WHS measure has to be rejected. The variables used are age, education and nationality of the wife, her attitude towards tradition, her non-wage income, the number of children, six dummies for the age of children living inside or outside the household, nationality of the mother, education of parents and social status of fathers' jobs (Treiman's prestige scale). The exogeneity of these variables was tested adding them to the probit of the husband's JEM and performing a simple significance test (based on the z-statistics).

Figure 4. Husbands' predicted probability to have a good job-education-match (GJEM=1)



Source: Swiss Household Panel (SHP) 1999

Table 7. Evaluation of the impact within-household specialisation (WHS) on the predicted probability of having a good job-education-match (JEM)

WHS: difference in housework hours:	0		15		30	
	dF/dx (2)	X (3)	dF/dx (2)	X (3)	dF/dx (2)	X (3)
Age	0.006 (0.95)	42.15 -	0.004 (0.95)	42.15 -	0.004 (0.95)	42.15 -
Experience	0.001 (0.63)	20.81 -	0.001 (0.63)	20.81 -	0.001 (0.63)	20.81 -
Part-time job	-0.004 (0.10)	0.00 -	-0.003 (0.10)	0.00 -	-0.002 (0.10)	0.00 -
Tenure	0.005 (3.09)***	11.21 -	0.004 (3.09)***	11.21 -	0.003 (3.09)***	11.21 -
Non-Swiss nationality	0.025 (0.73)	0.00 -	0.020 (0.73)	0.00 -	0.015 (0.73)	0.00 -
Difference in housework hours (Wife's - husband's hours)	0.002 (2.13)**	0.00 -	0.002 (2.13)**	15.00 -	0.001 (2.13)**	30.00 -
Husband's predicted probability of having a good JEM	0.893		0.920		0.942	
Number of observations	787		787		787	

Robust z-statistics in parentheses, **significant at 5%, *** significant at 1%

Source: Swiss Household Panel (SHP) 2000

Notes

- (1) GJEM is a dichotomous variable taking value 1 when a husband's educational qualifications correspond to his job and 0 if his qualifications are superior to his job.
- (2) Probit estimation, marginal effects.
- (3) The probability is evaluated for a Swiss man, married, working in a full-time job, having average age, experience and tenure, having obtained vocational training as the highest education degree. Furthermore he reports to have no difficulties in his professional or private life and his' satisfaction with health status and social skills are at average levels.

Table A1. Descriptive statistics of total sample and sample of matched spouses

<i>Mean (standard deviation)</i>	<i>Husband's characteristics:</i>		<i>Wife's characteristics:</i>	
	Total sample	Sample with matched spouses (3)	Total sample	Sample with matched spouses (3)
Good job-education-match (GJEM) (1)	0.878	0.865	-	-
Age	43.33 (10.40)	42.15 (10.34)	42.34 (11.76)	39.8 (10.20)
Market work experience (2)	22.34 (12.09)	20.81 (11.92)	6.75 (9.40)	7.53 (9.51)
Market work hours, weekly	45.05 (9.74)	44.14 (9.03)	15.23 (16.52)	16.84 (16.38)
Education years	13.67 (2.41)	13.80 (2.46)	12.45 (2.40)	12.79 (2.37)
Tenure	11.86 (10.61)	11.21 (10.32)	-	-
Non-Swiss nationality	0.201	0.182	0.193	0.167
Housework hours, weekly	6.10 (6.00)	6.10 (5.72)	20.12 (11.34)	19.77 (10.91)
Consensual partnership	0.091	0.087	0.091	0.087
Dependent child living in household	0.554	0.593	0.554	0.593
Number of observations	1507	787	1570	787

Source: Swiss Household Panel (SHP) 1999

Notes:

- (1) GJEM is a dichotomous variable taking value 1 when a husband's educational qualifications correspond to his job and 0 if his qualifications do not correspond to his job.
- (2) Years since last long interruption of market work (min. 6 month).
- (3) The sample of matched spouses includes only the observations of men and women clearly identified as a couple, and where for both spouses non-missing information on all the relevant variables is available.

Table A2. The impact of within-household specialisation (WHS) on husbands' job-education-match (JEM) – with controls for wives' potential earnings (1)

Probit, marginal effects are presented <i>Dependent variable:</i> Good job-education-match (GJEM) (2) (3)	All households	
	(A)	(B)
<i>Wife's education:</i>		
Compulsory schooling	-0.005 (0.10)	0.002 (0.04)
General training school	0.053 (0.67)	0.050 (0.59)
Vocational school, full time	0.036 (0.81)	0.048 (1.19)
Maturity	-0.043 (1.24)	-0.044 (1.08)
Vocational high education	0.000 (0.01)	-0.011 (0.15)
Technical or vocational school	-0.135 (1.16)	-0.140 (0.85)
Vocational high school	0.028 (0.49)	0.056 (0.88)
University, higher specialised school	-0.058 (1.30)	-0.108 (1.62)
<i>WHS:</i>		
Difference in housework hours (wife's - husband's hours, weekly)	0.002 (2.15)**	-
Difference in total housework hours, including child care (wife's - husband's hours, weekly)	-	0.002 (2.15)**
Number of observations (4)	787	439
Log-likelihood value	-278.2	-140.1
Pseudo <i>R</i> -squared	0.111	0.137

Robust z-statistics in parentheses, ** significant at 5%

Source: Swiss Household Panel (SHP) 1999

Notes:

- (1) Potential earnings of wives is approximated by their education measured in 9 levels. Therefore, 8 dummies are added to the probit equation, the default being a women having certified vocational training as her highest education degree.
- (2) GJEM is a dichotomous variable taking value 1 when a husband's educational qualifications correspond to his job and 0 if his qualifications are superior to his job.
- (3) Also included in the regression are the husband's education, age, age squared, market work experience, tenure, health, dummies for part-time working, for a Non-Swiss nationality, for difficulties in professional or private life and for persons being not married but living in a consensual partnership. Sample definition see Table 1.