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Forschungsinstitut zur Zukunft der Arbeit Institute for the Study of Labor

Intergenerational Transmission of Education: An Alert to Empirical Implementation

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ABSTRACT

Intergenerational Transmission of Education: An Alert to Empirical Implementation

The intergenerational transmission of education is certainly a problem that continues to challenge most countries. The level of education that an individual rises to is linked to the education level(s) of her/his parents. This note serves as an alert to researchers undertaking empirical investigation into how the parents' education should be considered with regard to the child's. Using Portuguese data we conclude that the parents should be viewed as a unit (i.e. as a couple), and we should examine all of the different education combinations, avoiding the temptation to aggregate them in larger categories.

JEL Classification: I21, J11

Keywords: transmission of education, human capital, parent's education

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Introduction

There is an extensive literature relating child's educational achievement to the education or income of their parents (Becker, 1988, Becker & Tomes, 1986, Haveman & Wolfe, 1995, Oosterbeek, 1995, Heineck & Riphahn, 2009, Pascual, 2009, Rumberger, 2009, just to cite a few). The model behind these studies is one where parents decide the allocation of resources to consumption and investment either on assets or human capital of their children. More education implies higher income and therefore a larger choice set allowing the choice of more human capital for their children. In this sense there is an intergenerational transmission of education meaning that children from parents with high education tend to attain high education while children from parents with low education tend to attain low education.

This conclusion has been tested and found to be valid in several empirical works. The issue I address in this note is the way that the education of the parents has been treated in some of these works. Some authors have considered the education of the parents as the highest level attained between the parents (e. g., Heineck & Riphahn, 2009), the highest level attained by both parents (e. g., Rumberger, 2009), or consider them separately (e. g., Pascual, 2009).

Using Portuguese data we test the following hypothesis:

- 1) Gender blindness the gender of the parent having the higher education is unimportant.
- 2) What counts is the highest level attained by at least one of the parents.
- 3) What counts is the highest level that both parents achieved.
- 4) The effect of both parents having education is equal to adding the separate effects.

Data and methods

We use IEFA¹ (Adult education and training survey -2007) data. Our data comprises 11,289 interviews (5,350 males, 5,939 females) in which the respondents were asked the educational level of the parents and their situation in the labor market while they were growing up (age 12 to 16).

In the dataset there are three educational levels for the parents from which the respondents could choose:

BAS – corresponding to less than or equal to 9 years of education;

SEC – corresponding to 11 or 12 years of education;

¹ This survey was carried out by Statistics Portugal and took place in all European Member States, following methodological guidelines issued by Eurostat.

HIG – Higher education degree,

We consider only those cases where we have information about both parents (10,436 observations). M stands for mother and F stands for father, so, as an example, MBAS_FBAS represents a couple in which both partners have BAS education.

In Table I we see the distribution of education among parents.

	N.	%
MBAS_FBAS	9,538	91.42
MBAS_FSEC	182	1.74
MBAS_FHIG	104	1.00
MSEC_FBAS	105	1.01
MSEC_FSEC	134	1.28
MSEC_FHIG	86	0.82
MHIG_FBAS	77	0.74
MHIG_FSEC	43	0.41
MHIG_FHIG	164	1.57
Total	10,433	100.00

Table I. Parents' educational achievement

We see that more than 90% of the individuals have both parents with at most a degree corresponding to 9 years of education.

In the dataset the education of the individual (child) appears in four categories, one more than those of the parents. The extra category is of individuals who attained no formal education. The categories and the distribution of education appear below.

NONE - no formal education;

BAS – corresponding to less than or equal to 9 years of education;

SEC – corresponding to 11 or 12 years of education;

HIG – Higher education degree;

=	N	
	N.	%
NONE	514	4.93
BAS	7,098	68.03
SEC	1,675	16.05
HIG	1,146	10.98
Total	10,433	100.00

Table II – Individual's education

We perform an ordered probit (four education levels). We use as explanatory variables parents education. We performed independent regressions depending on the gender of the individual.

Findings and conclusions

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We use as reference group an individual whose parents both have basic education.

For the sample of females we have the following results:

Educational	Coef.	Std. Err.	Z	P> z
level				
MBAS_FSEC	1.296761	.1227121	10.57	0.000
MBAS_FHIG	1.437740	.1614672	8.90	0.000
MSEC_FBAS	1.195232	.1446895	8.26	0.000
MSEC_FSEC	1.646253	.1451327	11.34	0.000
MSEC_FHIG	1.888777	.1911292	9.88	0.000
MHIG_FBAS	1.133614	.1866223	6.07	0.000
MHIG_FSEC	1.864989	.2755164	6.77	0.000
MHIG_FHIG	2.015568	.1390118	14.50	0.000
cut1	-1.522843	.0274417		
cut2	.6939479	.0192910		
cut3	1.3518650	.0242860		

And for males we have the following results:

Ordered probit regression	Number of obs $=$ 4958
	LR chi2(8) = 786.53
	Prob > chi2 = 0.0000
Log likelihood = -3928.0614	Pseudo R2 = 0.0910

	Coef.	Std. Err.	Z	P> z
MBAS_FSEC	1.252195	.1143128	10.95	0.000
MBAS_FHIG	1.459748	.1539760	9.48	0.000
MSEC_FBAS	1.539607	.1629626	9.45	0.000
MSEC_FSEC	1.513769	.1365484	11.09	0.000
MSEC_FHIG	1.73789	.1759110	9.88	0.000
MHIG_FBAS	1.544943	.1730826	8.93	0.000
MHIG_FSEC	2.044571	.2470063	8.28	0.000
MHIG_FHIG	1.943243	.1353325	14.36	0.000
cut1	-1.728752	.0332022		
cut2	.8654607	.0214239		
cut3	1.66038	.0298699		

Table IV – Males' education

Hypothesis 3) is rejected as having at least one parent with higher education has a positive effect that is significantly different from zero. The testing of the other hypothese appears in the table below.

Table V – Testing of hypothese				
	For the female sample		For the male sample	a 1 .
	Test	Conclusion	Test	Conclusion
Gender	MBAS_FSEC= MSEC_FBAS	Do not reject	MBAS_FSEC= MSEC_FBAS	Do not reject
blindness	chi2(1) = 0.29	hypothesis 1)	chi2(1) = 2.14	hypothesis 1)
	Prob > chi2 = 0.5885		Prob > chi2 = 0.1434	
	MBAS_FHIG= MSUP_FBAS	Do not reject	MBAS_FHIG= MSUP_FBAS	Do not reject
	chi2(1) = 1.54	hypothesis 1)	chi2(1) = 0.14	hypothesis 1)
	Prob > chi2 = 0.2148		Prob > chi2 = 0.7103	
	MHIG_FSEC=MSEC_FHIG	Do not reject	MHIG_FSEC=MSEC_FHIG	Do not reject
	chi2(1) = 0.01	hypothesis 1)	chi2(1) = 0.01	hypothesis 1)
	Prob > chi2 = 0.9432		Prob > chi2 = 0.9432	
Both parents =	MBAS_FSEC=MSEC_FSEC	Reject	MBAS_FSEC=MSEC_FSEC	Reject
at least one	chi2(1) = 3.46	hypothesis 2)	chi2(1) = 2.23	hypothesis 2)
parent	Prob > chi2 = 0.0629		Prob > chi2 = 0.1353	
	MSEC_FBAS= MSEC_FSEC	Reject	MSEC_FBAS= MSEC_FSEC	Reject
	chi2(1) = 4.94	hypothesis 2)	chi2(1) = 0.02	hypothesis 2)
	Prob > chi2 = 0.0263		Prob > chi2 = 0.9021	
	MHIG_FBAS= MHIG_FHIG	Reject	MHIG_FBAS=MHIG_FHIG	Reject
	chi2(1) = 14.59	hypothesis 2)	chi2(1) = 3.37	hypothesis 2)
	Prob > chi2 = 0.0001		Prob > chi2 = 0.0665	
	MBAS_FHIG= MHIG_FHIG	Reject	MBAS_FHIG=MHIG_FHIG	Reject
	chi2(1) = 7.50	hypothesis 2)	chi2(1) = 5.72	hypothesis 2)
	Prob > chi2 = 0.0062		Prob > chi2 = 0.0168	
Both	MSEC_FSEC=MSEC_FBAS+	Reject	MSEC_FSEC=MSEC_FBAS+	Reject
parents=parent	MBAS_FSEC	hypothesis 4)	MBAS_FSEC	hypothesis 4)
A+parent B	chi2(1) = 12.73		chi2(1) = 28.57	
	Prob > chi2 = 0.0004		Prob > chi2 = 0.0000	
	MHIG_FHIG=MHIG_FBAS+	Reject	MHIG_FHIG=MHIG_FBAS+	Reject
	MBAS_FHIG	hypothesis 4)	MBAS_FHIG	hypothesis 4)
	chi2(1) = 3.90		chi2(1) = 15.93	
	Prob > chi2 = 0.0484		Prob > chi2 = 0.0001	

Table V – Testing of hypothese

Given the above findings² we conclude that in the empirical work we should consider the different pairs of possibilities of education of the parents and test if we can join some of them in larger categories. In the Portuguese case we cannot.

Gender blindness seems to be the exception, as we could not statistically reject the hypothesis.

Finally, the effect of the parents' education is not the same as the effect of the education of each parent added together. As a result, we should not treat them independently.

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² Similar results were obtained in a regression where we considered as explanatory variables not only the education of the parents but also their situation in the labour market and the age of the individual.

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