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## The Opt-Out Revolution: A Descriptive Analysis

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Discussion Paper No. 5089  
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## **ABSTRACT**

### **The Opt-Out Revolution: A Descriptive Analysis**

Using data from the 1980, 1990, and 2000 U.S. Census, I find little support for the opt-out revolution – highly educated women, relative to their less educated counterparts, are exiting the labor force to care for their families at higher rates today than in earlier time periods – if one focuses solely on the decision to work a positive number of hours irrespective of marital status or race. If one, however, focuses on both the decision to work a positive number of hours as well as the decision to adjust annual hours of work (conditional on working), I find some evidence of the opt-out revolution, particularly among white college educated married women in male dominated occupations.

JEL Classification: J13, J15, J16, J22

Keywords: opting out, female labor supply, extensive/intensive margin, race/ethnicity

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## 1. Introduction

A number of recent studies have attempted to find evidence to support the opt-out revolution—highly educated women, relative to their less educated counterparts, are exiting the labor force to care for their families at higher rates today than in earlier time periods—which was first sensationalized by Lisa Belkin in a 2003 *New York Times* article.<sup>1</sup> The preponderance of these studies focus on the employment side of opting-out (i.e., conditional on having children women exit the labor market)<sup>2</sup> and generally conclude that there is no evidence to support the popular press’ notion of an opt-out revolution, that is, the child penalty is not found to be higher for highly educated women compared to less educated women in recent years (see for example, Boushey 2005, Goldin 2006, Vere 2007, Cohany and Sok 2007, Fortin 2008, and Percheski 2008).<sup>3</sup>

What can account for the discrepancy in the press’ accounts of the opt-out revolution and the studies of the employment side of the opt-out revolution (henceforth referred to as the opt-out literature)? One possibility is the opt-out literature tends to focus on the child penalty unadjusted for demographic characteristics, such as, marital status and race/ethnicity (henceforth referred to

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<sup>1</sup> Many other media reports followed (see for example, Wallis 2004, Story 2005, and Trunk 2005).

<sup>2</sup> Several studies examine the fertility side of opting out. The results are mixed; Vere (2007) and Shang and Weinberg (2009) find increased fertility among college educated women while Percheski (2008) finds no change in fertility among professional women. Others have examined the impact of opting-out on labor market outcomes among alumni from top-ranked colleges. Goldin and Katz (2008) find no evidence of the opt-out revolution among Harvard University alumnae but do find that opting out impacts earnings, although differentially by occupation. Herr and Wolfram (2009) find that alumnae from Harvard University are less likely to stay attached to the labor market if their workplace lacks family friendly policies. Bertrand, Goldin, and Katz (2009) find that the gender wage gap among MBA graduates from the University of Chicago’s Booth School of Business can largely be explained by career interruptions from childbirth. Several studies also seek to understand why highly educated women are exiting the labor market and what the associated costs of exiting may be (see for example, Stone and Lovejoy 2004, Hewlett *et al.* 2005, Williams *et al.* 2006, Hewlett 2007, and Stone 2008); these studies tend to be more qualitative in nature.

<sup>3</sup> The one exception is Vere (2007) who does find evidence that highly educated women with children are exiting the labor market at higher rates than their less educated counterparts. However, he only focuses on women 27 years of age which is very young for childbearing, particularly in later cohorts for highly educated women.

as race).<sup>4</sup> The labor supply decisions of prime-aged women (henceforth referred to as women) with children who are not married are likely to be very different than those of their married women counterparts, particularly at the top of the education distribution, as these women are likely married to men who are also highly educated and have high earning power, affording them the ability to exit the labor market when they have children (see for example, Williams *et al.* 2006).<sup>5</sup> Indeed highly educated married women tend to be the focus of the media accounts (Williams *et al.* 2006). In addition, in 2000 marriage rates among women are higher at the top of the education distribution than at the bottom of the education distribution; this is in stark contrast to 1980 where highly educated women are generally less likely to get married than their less educated counterparts.<sup>6</sup> Marriage rates have fallen over time as well (see Table 1). Thus by combining married and non-married women, one may be underestimating the child penalty, particularly in recent years.

Similarly, the labor supply decisions of non-Hispanic white (henceforth referred to as white) women with children are likely to be very different than those of their non-white counterparts. The media attention is generally more focused on white affluent women in white collar jobs (Williams *et al.* 2006). Furthermore, white women are generally more attached to the

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<sup>4</sup> Boushey (2005) does allow for intercept effects of race/ethnicity and marital status when estimating the effect of children on women's labor force attachment but does not allow for slope effects (i.e., she does not interact marital status or race/ethnicity with children). Cohany and Sok (2007) do report labor force participation by race, however they only do so for married women with infants. While Fortin (2008) estimates the effect of children on women's labor force participation separately by marital status and education, she only stratifies by two education categories: less than college and college graduates. Fortin (2008) also analyzes the effect of children on black women's labor force participation, however, she does not allow for differential effects of children by education or marital status for this group of women. Another demographic characteristic is immigrant status. Results that focus solely on non-immigrants are generally similar and available upon request.

<sup>5</sup> Cha (2009) finds that professional women with children, married to professional men, were substantially more likely to exit the labor market if their partner worked extreme hours (>60 hours) relative to their childless counterparts whose partners did not work extreme hours.

<sup>6</sup> The marriage rates for women aged 25-44 in 1980 are 69% (less than high school), 76% (high school), 72% (some college) and 69% (college). The marriage rates for women aged 25-44 in 1990 are 58% (less than high school), 68% (high school), 66% (some college), and 66% (college). The marriage rates for women aged 25-44 in 2000 are 57% (less than high school), 63% (high school), 63% (some college), and 66% (college).

labor market, are more likely to have a college education, more likely to be married, and are less likely to have children, and if they do have kids, they have fewer of them, relative to their non-white counterparts (see Table 1). Moreover, racial differences in educational attainment have become more pronounced over time. All of these racial differences suggest that the child penalty may differ substantially by race, and these differences may shed light on the discrepancies between the results presented in the media and the opt-out literature.

Another possibility is the opt-out literature generally focuses on one aspect of female labor supply, the decision to work a positive number of hours (i.e., the extensive margin), as opposed to conditional on participating in the labor market, the decision to adjust annual hours of work (i.e., the intensive margin).<sup>7</sup> I argue focusing on both margins may be important because women, particularly highly educated women, may not be choosing to exit the labor market entirely, but may be opting for reduced hours of work. There is support for this in the qualitative literature on opting out (see for example, Hewlett *et al.* 2005, Hewlett 2007, and Stone 2008), that is, highly educated women are seeking alternative occupations which are less demanding, afford them more flexibility, and require a reduced time commitment. Moreover, Belkin (2003) recognized this possibility in her *New York Times Magazine* article. This then suggests that defining opting out simply on a woman's decision to exit the labor market may be masking some of the employment aspect of the opt-out revolution.

A third possibility is defining highly educated women simply as college educated may obscure the opt-out phenomenon. It may be that it is college educated women in specific managerial/professional occupations, particularly male dominated occupations (e.g., architects, engineers, lawyers, and physicians), that are more likely to opt out of the labor market in terms

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<sup>7</sup> Percheski (2008) does consider alternative measures of labor supply (full-time year round employment and works 50+ hours per week) but does not try to disentangle the relative roles of extensive and intensive margin.

of both the extensive and intensive margin.<sup>8</sup> On the one hand, women in these occupations may be more inclined to exit the labor market entirely because these occupations offer the least amount of job flexibility (see Stone 2008). On the other hand, women in these types of occupations likely worked extreme hours (e.g., 60+ hours/week) prior to having their child/children, and may be able to adjust their hours down to accommodate their new family roles. In addition, it is this group of women that the media tends to focus on (see Williams et al., 2006).

In this paper, I show that, as in the existing opt-out literature, there is little evidence to support the opt-out revolution if one focuses solely on the extensive margin. Specifically, using the 1980, 1990, and 2000 U.S. Census, I find that the child penalty for women fell substantially between 1980 and 1990 irrespective of one's level of education. Moreover, the child penalty for this group did not increase in recent years; in fact the child penalty continued to fall (although to a lesser degree) between 1990 and 2000 for all education categories except for the college educated category where the child penalty in 2000 remained roughly equal to the child penalty in 1990. The same patterns generally hold for college educated married women irrespective of race, that is, the child penalty did not increase between 1990 and 2000.

As predicted, I find that the overall child penalty in female labor supply is higher if one focuses on both the extensive and intensive margin, suggesting the importance of accounting for both margins. Despite this, the overall child penalty in female labor supply for college educated married women does not rise, as predicted by the opt-out revolution, but falls in recent years. While this holds true for white and Hispanic college educated married women, the reverse is true

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<sup>8</sup> While Percheski (2008) does examine women in managerial/professional occupations, as well as the subset of traditionally male and female dominated occupations within this occupational category, she does not examine detailed occupations in managerial/professional occupations, traditionally male occupations, or traditionally female occupations. Focusing on aggregate categories may mask the opt-out phenomenon as there are different degrees of flexibility and the incidence of extreme hours in the detailed occupations within each occupation category.



for their non-Hispanic black (henceforth referred to as black) and Asian counterparts. Strikingly, I find that the child penalty does increase between 1990 and 2000 for white college educated married women who are in male dominated occupations whereas for white college educated married women in female dominated and mixed occupations the child penalty decreases between 1990 and 2000. Taken together, these patterns suggest that there is some support for the opt-out revolution particularly among the group of women that the media tends to focus on, that is, white college educated married women who are in male dominated occupations.

The remainder of the paper is as follows. The data is discussed in Section 2. Section 3 examines labor force attachment by marital status, race, year and level of education. Section 4 presents the results for the child penalty decomposition. Section 5 presents an in depth analysis for white college educated married women. The conclusions are presented in Section 6.

## **2. Data**

I use data from the *Integrated Public Use Microdata Series* (IPUMS) 5% sample of the 1980, 1990, and 2000 United States Census. This data is ideal for my purposes because it has detailed information on family characteristics (marital status and presence of children), labor market characteristics (usual hours worked per week, weeks worked in the past calendar year, and occupation), and demographic characteristics (age and education), as well as having large samples of women within educational, racial/ethnic, and occupational categories.<sup>9</sup>

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<sup>9</sup> Ideally, one would supplement the Census data with data from the Current Population Survey (CPS) to get information on the years between the censuses, as well as information beyond 2000. Unfortunately, due to small sample sizes this is not possible for all educational, racial/ethnic, and occupational categories. Despite this shortcoming, I am able to use the CPS to investigate the opt-out phenomenon for Non-Hispanic white college educated married women. This analysis is presented in detail in Section 5.

I restrict the sample to women between the ages of 25-44 who are employed in the civilian labor force and do not have allocated values on any of the variables of interest.<sup>10</sup> I focus on 25-44 year olds to ensure that women have completed their formal education, as well as allowing them sufficient time to begin forming their families. I only focus on the civilian labor force as the armed forces labor force is very different and likely attracts women with different preferences for work and family.

I consider four racial groups, white, black, Hispanic, and Asian. These are constructed as indicator variables equal to one if the respondent is of that particular race, and zero otherwise.<sup>11</sup> Table 1 presents summary statistics by race and year. The proportion of the female population that is white (non-white) decreased (increased) over the time period, with the biggest increase realized by Hispanics. Whites (Hispanics) accounted for 80.7 (6.1), 76.5 (8.1), and 70.7 (10.9) percent of the female population in 1980, 1990, and 2000, respectively.

I consider three measures of labor supply. The first measure of labor supply considered is unconditional annual hours of work. This is defined as usual hours per week times weeks worked in the past calendar year. It therefore includes women who are not working as by construction it includes zero values. I also consider conditional annual hours of work measured as annual hours of work conditional on working positive annual hours of work. Finally, I consider labor force attachment (LFA) which is defined to equal one if a respondent worked positive annual hours, and zero otherwise.<sup>12</sup> By focusing on these three measures, I am able to examine both the intensive and extensive margins (discussed in detail below).

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<sup>10</sup> Specifically, I drop an individual if they have an allocated value for age, sex, race, Hispanic, education, usual hours of work, weeks worked per year, employment status, marital status, and relationship to household head.

<sup>11</sup> In order to construct these variables I rely on two IPUMS variables: race (RACED) and Hispanic origin (HISPAND).

<sup>12</sup> Alternatively I could have used labor force participation (LFP) which equals one if a respondent was employed or actively seeking employment in the reference week, and zero otherwise. While this is the measure generally considered in the opt-out literature, for my purposes labor force attachment (LFA) is needed to determine the

Women were more attached to the labor market in 2000 relative to 1980 irrespective of the labor supply measure considered, although the biggest increase in labor supply occurred between 1980 and 1990. For example, LFA rates in 1980, 1990, and 2000 are 68.3, 77.2, and 78.7 percent, respectively. The same general pattern holds for all races, although it is more pronounced for white women. Focusing again on LFA, white women's LFA rates increased 12.3 percentage points between 1980 and 2000 while black, Hispanic, and Asian women's LFA rates increased by 9.8, 8.1, and 2.3 percentage points, respectively, in this time period.

To determine the child penalty, information on presence of children in the household is needed. I use the family relationship variables constructed in the IPUMS data to ascertain the presence of children. The first indicator variable is constructed based on the number of own children (biological, adopted, or step children) in the household (NCHILD). Specifically, I create an indicator variable equal to 1 if there are no own children in the household, and zero otherwise. A second indicator variable is constructed based on the number of own children under age 5 in the household (NCHLT5). Specifically, I create an indicator variable equal to 1 if there are any own children less than 5 in the household, and zero otherwise. For prime-aged women, own children in the household is a relatively good approximation for fertility (see Rindfuss 1976). The number of own children decreased from 1.668 in 1980, to 1.399 in 1990, and 1.386 in 2000 while the proportion of women who had no children increased from 25.6 percent, to 31.6 percent in 1990, and 33.0 percent in 2000. The proportion of women having children less than five is relatively stable between 1980 and 2000; 27.6, 27.2, and 26.5 percent in 1980, 1990, and 2000, respectively. While the levels are different across racial groups (e.g.,

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relative roles of the intensive and extensive margins of labor supply. Results based on LFP are very similar to results based on LFA and are available upon request.

non-whites have more children than their white counterparts), the overall patterns tend to hold for all racial groups.

I also create a marital status indicator variable equal to one for married, spouse present, and zero otherwise. Marriage rates have fallen over time from 72.9 percent in 1980, to 65.6 percent in 1990, and to 63.4 percent in 2000. Blacks experienced the largest decline in marriage rates between 1980 and 2000 (11.8 percentage points) followed by whites (8.6 percentage points), Asians (6.8 percentage points), and Hispanics (6.8 percentage points).

In addition, I consider 4 educational categories, less than high school (grade 11 or less), high school (grade 12), some college (1-3 years of college), and college (4+ years of college).<sup>13</sup> These are constructed as indicator variables equal to one if the respondent has attained that particular level of education, and zero otherwise. Women have acquired more education between 1980 and 2000; 18.0 (18.5), 10.9 (23.7), and 8.8 (29.6) percent of women had a less than high school (college) degree in 1980, 1990, and 2000, respectively. The increase in college attendance between 1980 and 2000 is highest for Asians (37.2 to 51.5 percent), followed by whites (19.9 to 32.7 percent), blacks (11.1 to 19.1 percent), and Hispanics (7.4 to 13.7 percent).

Finally, I examine detailed occupations that fall into 3 occupational categories within the managerial/professional occupations: male dominated (roughly 70 percent male or higher in 1980, 1990 and 2000), female dominated (roughly 30 percent male or lower in 1980, 1990 and 2000), and mixed occupations (roughly between 40 and 60 percent male in 1980, 1990 and 2000). These categories are constructed based on the “occupation, 1990 basis” (OCC1990)

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<sup>13</sup> These categories are constructed based on the educational attainment recode variable (EDUREC) in the IPUMS data which combines two separate IPUMS variables HIGRADE and EDUC99. The former is available from 1940 to 1980 and the latter is available from 1990-2000. The recoded variable was constructed to allow for analysis across the years. The main differences are high school based on EDUC99 is 12<sup>th</sup> grade no diploma and high school diploma or GED, some college based on EDUC99 is some college no degree, occupational associate degree, and academic associate degree, and college based EDUC99 is bachelor’s degree, master’s degree, professional degree and doctorate degree.

variable in the IPUMS data for individuals between the ages of 18 to 64. An ideal feature of the IPUMS occupation question is that it applies to individuals who had worked within the previous five years which allows one to examine both the extensive margin and intensive margin. Moreover, the 5 year time frame coincides with the presence of young children, i.e., children less than 5 years of age. See Appendix Table 2 for a detailed list of occupations by year that fall under each occupation category, as well as the percent male within each detailed occupation category.

Unfortunately due to small sample sizes (particularly in male dominated occupations), I am only able to consider a subset of detailed occupations within each occupation type (see Appendix Table 2 for the sample size of white college educated married women by occupation, year, and presence of children).<sup>14</sup> The detailed male dominated occupations considered are architects, electrical engineers, physicians, lawyers, clergy/religious workers. The detailed female dominated occupations considered are registered nurses, dietitians/nutritionists, occupational therapists, primary school teachers, and librarians. The mixed occupations considered are human resources and labor relations manager (henceforth referred to as human resources), accountants and auditors, secondary school teachers, economists and market researchers (henceforth referred to as economists), and editors/reporters.

### **3. Labor Force Attachment (LFA) by Year, Education, and Race**

Following the opt-out literature, I first focus on the extensive margin (i.e., the decision to work a positive number of hours). Figure 1 presents the labor force attachment (LFA) rates for all women aged 25-44 by presence of young children, year, and education. Figure 1 also presents the child penalty measured as:

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<sup>14</sup> An occupation is considered if there are at least 25 women without kids/with kids less than 5 in each year.

$$\frac{\overline{LFA}_{NK} - \overline{LFA}_{KLT5}}{\overline{LFA}_{NK}} \quad (1)$$

where  $\overline{LFA}$  is mean labor force attachment and NK and KLT5 denote women without children and women with young children, respectively.<sup>15</sup>

There are several noteworthy patterns. First, LFA of women with young children has increased over time irrespective of education category, although the increase is more pronounced between 1980 and 1990 than between 1990 and 2000. For example, the LFA rate for women with young children increased from 52.0 in 1980 to 65.0 percent in 1990 and 68.0 percent in 2000. Perhaps surprisingly, the same is not true for women without children. For less educated (high school or below) the LFA of childless women decreased over time (e.g., 56.5, 54.8 and 52.4 percent in 1980, 1990, and 2000 for women with less than high school) while for more educated (some college and above) the LFA of childless women remained relatively flat between 1980 and 2000 (e.g., 94.9, 95.6 and 94.4 percent in 1980, 1990, and 2000 for women with college).

Second, LFA is higher among women with higher levels of education relative to their less educated counterparts irrespective of year or presence of children. Focusing on 2000, it can be seen that the LFA rate for childless women with less than high school is 52.4 percent while their high school, some college, and college counterparts have LFA rates of 77.7, 89.5, and 94.4 percent, respectively.

Third, childless women have higher LFA rates than women with young children. This is again true regardless of the education category or year considered. For example, in 1990 the LFA rate for childless women is 77.6 percent while the LFA rate for women with young children LFA rate is 64.5 percent.

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<sup>15</sup> Percheski (2008) measures the child penalty analogously.

Fourth, there is no clear pattern by education in the child penalty in a given year. In 1980 and 1990 the child penalty rises (falls) as one acquires more education at the bottom (top) of the education distribution while in 2000 the child penalty generally rises with education. For example, in 1980 (2000) the child penalty for women with less than high school, high school, some college, and college is 31.7 (16.6), 40.7 (18.9), 37.5 (18.9), and 34.2 (21.8), respectively. There is, however, a clear pattern in the child penalty over time in a given education category. Specifically, the child penalty fell between 1980 and 1990 for all education groups, while between 1990 and 2000 it fell (although a much smaller decline than in the previous time period) for all education groups except college educated women where the penalty remained relatively constant. For instance, the child penalty for high school educated women went from 40.6 percent in 1980, to 26.6 percent in 1990, to 18.9 percent in 2000 while for college educated women the child penalty went from 34.2 percent in 1980, to 22.2 percent in 1990, to 21.8 percent in 2000. Taken together, these results provide little support for the media accounts of the opt-out revolution.<sup>16</sup>

Figure 1 combines all women irrespective of marital status. As previously discussed, the media tends to focus on highly educated married women who may have the luxury to opt-out of the labor market because they tend to be married to highly educated men with high earning potential (see for example, Williams *et al.* 2006). Moreover, marriage rates for college educated women in 2000 are higher than those of their less educated counterparts; this was not true in 1980. Marriage rates have fallen over time as well (see Table 1), although less so for college educated women. By combining married women with non-married women, one may be underestimating the child penalty, particularly in recent years. Thus, for the remainder of the

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<sup>16</sup> The same patterns are found if one examines the child penalty by education, age, and year, as well as by education, age, and birth cohort (results available upon request). This is discussed in greater detail for a subset of the sample in Section 5.

analysis I focus on married women. Figure 2 presents the LFA rates and the child penalty for married women by presence of young children, education, and year. While there is evidence that the child penalty is underestimated in recent years (particularly at the bottom of the education distribution), it remains the case that the child penalty did not rise between 1990 and 2000 for college educated women.

Allowing for differential effects by race may help shed light on the discrepancy between the media accounts of the-opt out revolution and the results presented to this point. Specifically, it is unlikely that the child penalty will be the same by race given there are substantial racial differences in labor force attachment, educational attainment, marital status, and presence of children (see Table 1). Figures 3, 4, 5, and 6 present the LFA rates and the child penalty for married women by presence of young children, education, and year for whites, blacks, Hispanics, and Asians, respectively.

Not surprisingly, the results for married white women (Figure 3) mirror those presented for all married women (Figure 2) given white women make up such a large share of the total population. The patterns for non-white married women differ somewhat from their white counterparts. In particular, the child penalty generally falls as non-whites acquire more education irrespective of year.<sup>17</sup> The child penalty is generally smaller in magnitude for non-whites, and this is particularly true for married black women.<sup>18</sup> For example, the child penalty for white married women in 2000 is 29, 26, 24, and 24 percent for less than high school, high school, some college, and college, respectively (see Figure 3) while the child penalty for black married women in 2000 is 1, 3, 7, and 7 percent for less than high school, high school, some

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<sup>17</sup> The main exception is for Black married college educated women in 2000 who face a higher child penalty than their less educated counterparts.

<sup>18</sup> The main exception is for Asian married women with less than high school in 1990 who face a child penalty of 42 percent (see Figure 6) while their white married counterparts face a child penalty of 28 percent (see Figure 3).



college, and college, respectively (see Figure 4). What is true for all racial groups, however, is that the child penalty for college educated women did not increase between 1990 and 2000.

The LFA patterns revealed in Figures 1 through 6 are consistent with the evidence from the opt-out literature (see Boushey 2005, Cohany and Sok 2007, and Percheski 2008) and suggest that the opt-out revolution, as sensationalized by Lisa Belkin in a 2003 *New York Times* article, is not generally supported in the data. In other words, the child penalty did not increase in recent year for college educated women, nor did college educated women face a higher child penalty relative to their less educated counterparts within a given year, irrespective of race or marital status.<sup>19</sup>

This conclusion however is based on the extensive margin (i.e., the decision to work a positive number of hours), and does not consider the intensive margin (i.e., conditional on participating in the labor market, the decision to adjust annual hours of work). I argue this is important because women, particularly highly educated women, may not be choosing to exit the labor market entirely, but may be opting for reduced hours of work. This then suggests that defining opting out as simply labor market attachment may be masking some of the “opt-out revolution”. The remainder of the paper examines the child penalty based on both the extensive and intensive margin.

#### **4. Decomposition of the Overall Child Penalty by Year, Education, and Race**

To formally analyze both the extensive and intensive margin, I first note that mean unconditional annual hours is equal to conditional annual hours times LFA. Given this, the child

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<sup>19</sup> The main exception is for black college educated married women in 2000 who face a child penalty of 7 percent, while black married women with less than high school, high school, and some college face a child penalty of 1, 3, and 7 percent, respectively (see Figure 4).

penalty in unconditional annual hours (henceforth referred to as the overall child penalty) can be written as

$$\frac{\overline{UAH}_{NK}^M - \overline{UAH}_{KLT5}^M}{\overline{UAH}_{NK}^M} = \frac{\overline{CAH}_{NK}^M * \overline{LFA}_{NK}^M - \overline{CAH}_{KLT5}^M * \overline{LFA}_{KLT5}^M}{\overline{UAH}_{NK}^M} \quad (2)$$

where  $\overline{UAH}$  is mean unconditional annual hours,  $\overline{CAH}$  is mean conditional annual hours,  $\overline{LFA}$  is mean labor force attachment, and M, NK, and KLT5 denote married, women without children, and women with young children, respectively. I now add and subtract from the numerator of the right hand side (RHS) of equation (2) the counterfactual mean unconditional annual hours for married women without children if they had the mean labor force attachment of married women with young children ( $\overline{CAH}_{NK}^M * \overline{LFA}_{KLT5}^M$ ):

$$\frac{\overline{UAH}_{NK}^M - \overline{UAH}_{KLT5}^M}{\overline{UAH}_{NK}^M} = \frac{\overline{CAH}_{NK}^M * \overline{LA}_{NK}^M - \overline{CAH}_{NK}^M * \overline{LA}_{KLT5}^M + \overline{CAH}_{NK}^M * \overline{LA}_{KLT5}^M - \overline{CAH}_{KLT5}^M * \overline{LA}_{KLT5}^M}{\overline{UAH}_{NK}^M} \quad (3)$$

Collecting terms:

$$\frac{\overline{UAH}_{NK}^M - \overline{UAH}_{KLT5}^M}{\overline{UAH}_{NK}^M} = \frac{(\overline{LA}_{NK}^M - \overline{LA}_{KLT5}^M) * \overline{CAH}_{NK}^M}{\overline{UAH}_{NK}^M} + \frac{(\overline{CAH}_{NK}^M - \overline{CAH}_{KLT5}^M) * \overline{LA}_{KLT5}^M}{\overline{UAH}_{NK}^M} \quad (4)$$

The first term on the RHS of equation (4) represents the portion of the overall child penalty due to the extensive margin and the second term represents the portion of the overall child penalty due to the intensive margin.<sup>20</sup> Dividing each term on the RHS of equation (4) by the overall

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<sup>20</sup> Instead, I could have added and subtracted from equation (3) the counterfactual mean unconditional annual hours for married women without children if they had the mean conditional annual hours of married women with young children ( $\overline{CAH}_{KLT5}^M * \overline{LA}_{NK}^M$ ). This adjusts the intensive margin, as opposed to the extensive margin (as is done in equation 3). While this yields a slightly larger role for the intensive margin, the overall patterns are the same (results available upon request).

child penalty (and multiplying by 100) gives the percent explained by the extensive and intensive margin, respectively.<sup>21</sup>

Before focusing on the decomposition results, it is worthwhile to discuss the patterns in female labor supply based on unconditional annual hours, as well as the patterns in the overall child penalty. Unconditional annual hours have generally increased for married women irrespective of presence of children or level of education (see Figure 7).<sup>22</sup> For example, college educated childless women worked 1624 unconditional annual hours in 1980 compared to 1835 and 1887 unconditional annual hours in 1990 and 2000, respectively. While this rise in unconditional annual hours holds for white married women (see Figure 8), it does not for their non-white counterparts (see Figures 9-11). For example, black married childless women with lower levels of education (high school or less) experience a fall in unconditional annual hours between 1990 and 2000 (see Figure 9). In addition, the overall child penalty is higher than the child penalty in LFA irrespective of race, education level, and year, suggesting the importance of accounting for both the extensive and intensive margin. For example, for college educated married women in 2000 the overall child penalty is 38.6 (see Figure 7) while the child penalty in LFA is 21.8 (see Figure 1).

Turning to the decomposition results (see Table 2), I find the intensive margin explains a larger share of the overall child penalty as married women acquire more education irrespective of the year considered. For example, in 2000 the intensive margin explains 26.6, 33.9, 40.9, and 42.8 percent of the overall child penalty for less than high school, high school, some college, and college, respectively (see column 3). Interestingly, the portion of the overall child penalty

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<sup>21</sup> Antecol and Steinberger (2009) use a similar technique to examine differences in labor supply between married women and partnered lesbian women.

<sup>22</sup> The main exception is for married women with young children who have less than high school between 1990 and 2000 where unconditional annual hours marginally fell.

explained by the extensive margin decreased between 1980 and 1990 and increased between 1990 and 2000 for all education categories, although it is more pronounced for college educated women. This suggests that married women, particularly college educated married women, in recent years are reverting back to adjusting their labor supply on the extensive margin as opposed to the intensive margin. This is noteworthy in its own right as it may reflect the fact that college educated married women may have less of an opportunity to adjust their hours of work any further due to workplace hour constraints. If they require more flexibility than their workplace can provide, then they may feel they are left with no choice but to exit the labor market. This phenomenon has been documented in the qualitative literature (see for example, Stone 2008).

The overall patterns hide some important racial differences. While similar results are found for white and Hispanic married women, the same is not true for black and Asian married women. For black married women, the intensive margin plays the biggest role for college educated women relative to their less educated counterparts in 1980 and 1990 (see columns 7 and 8) while in 2000 it is least important for college educated women relative to their less educated counterparts (see column 9). This is largely explained by the fact that the LFA rates of black married women with less than high school in 2000 were virtually identical, not because black married women with young children substantially increased their LFA between 1990 and 2000 in this education category but because childless black married women with less than high school substantially reduced their LFA during this time period (see Figure 4). There are no clear patterns in the importance of the intensive margin by level of education for Asian married

women.<sup>23</sup> For black (Asian) married college educated women, on the other hand, the relative role of the extensive margin remained roughly the same (decreased) between 1990 and 2000.

Despite the fact that the aforementioned patterns illustrate the importance of accounting for both the extensive and intensive margins, as with the child penalty in LFA, the overall child penalty does not increase in recent years for married college educated women (see Figure 7). This, however, masks important racial differences. Specifically, the child penalty for black (Asian) college educated married women increases by 1.5 (2.3) percentage points between 1990 and 2000 (see Figure 9 (11)). For white and Hispanic college educated married women, the child penalty falls by 3.6 and 3.1 percentage points, respectively, between 1990 and 2000 (see Figures 8 and 10, respectively). These patterns continue to suggest limited evidence in support of the opt-out revolution, particularly for white college educated married women who tend to be the focus of the media. This may be driven by the fact that college education is too coarse of a definition for this group. Thus the remainder of the paper presents an in depth examination of white college educated married women.

## **5. A Further Investigation of Non-Hispanic White College Educated Married Women**

The opt-out phenomenon among non-Hispanic white (henceforth referred to as white) college educated married women may be obscured for a number of reasons. First, it may be that limiting the analysis to 2000 is too early a time period in which one could detect the opt-out phenomenon, which was not brought to the public's attention until Belkin's 2003 *New York Times* article. Second, it may be that the propensity to opt-out of the labor market varies by age

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<sup>23</sup> According to Table 2 (columns 13-15), the relative importance of the intensive margin in 1980 is higher for Asian married women with some college, followed by less than high school, college, and high school while in 1990 (2000) the relative importance of the intensive margin is higher for Asian married women with some college (less than high school) followed by high school (college), college (some college), and less than high school (high school).

for college educated women as there may be large differences in the work behavior of young women relative to older women due to, for example, a change in preferences towards work and family (i.e., culture). Finally, it may also be that college educated women across occupation categories have different propensities to opt-out of the labor market as the opportunities for work-place flexibility vary by occupation. Each of these possibilities is examined in turn.

Table 3 presents the decomposition of the overall child penalty for white college educated married women by year: 1980, 1981-1983, 1984-1986, 1987-1989, 1990, 1991-1993, 1994-1996, 1997-1999, 2000, 2001-2003, 2004-2006, 2007-2009. The data for 1980, 1990, and 2000 continue to be based on the U.S. census data, while the data for the intervening years are from *Integrated Public Use Microdata Series (IPUMS) March Current Population Survey (CPS)* samples. These data are ideal for these purposes because IPUMS constructs the variables to align with those in the census data, however, the CPS sample sizes are small.<sup>24</sup> Given the small sample sizes, I pool over three year periods in the CPS. Interestingly, it can be seen that, if anything, evidence of the opt-out phenomenon actually occurs between the latter half of the 1990s and the early 2000s, as opposed to the mid- to the late- 2000s.<sup>25</sup> Specifically, the overall child penalty falls continuously from 60.3 in 1980 to 40.2 in 1997-1999, increases to 41.6 in 2000 and 42.2 in 2001-2003,<sup>26</sup> then falls thereafter (39.9 in 2004-2006 and 38.3 in 2007-2009).<sup>27</sup>

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<sup>24</sup> Ideally one would use this data throughout the analysis, however due to small sample sizes this is not possible for all educational, racial/ethnic, and occupational categories.

<sup>25</sup> Interestingly, a similar pattern is found if one solely examines the extensive margin. In particular, the child penalty in labor force attachment continuously falls from 37.1 in 1980 to 20.1 in 1997-1999, increases to 23.6 in 2000, and then falls thereafter (23.5 in 2001-2003, 22.6 in 2004-2006, and 21.1 in 2007-2009).

<sup>26</sup> If one pools the CPS over 1999-2001 in an attempt to get an approximation of the 2000 Census number (41.6), the child penalty is 42.0, which is very much in line with the Census data.

<sup>27</sup> This is generally true for all levels of education, although it is noisier in the 1990's for white married women with less than high school due to small cell sizes. Moreover, this holds true for white married college educated women across the age distribution.

Table 4 presents the decomposition of the overall child penalty for white college educated married women by year and 4 age groups, 25-29, 30-34, 35-39, and 40-44.<sup>28</sup> Not surprisingly the overall child penalty is smaller for the youngest women (25-29) and steadily increases until the oldest age group (40-44) where it begins to decline irrespective of year. For example, in 2000 the overall child penalty is 34.12, 42.77, 45.47, and 41.89 for white college educated married women aged 25-29, 30-34, 35-39, and 40-44, respectively. Moreover, the intensive margin is of greater importance for young women relative to their older counterparts for all years. This lends support to the hypothesis that work behavior differs by age. Despite these age differences, there continues to be little support for the opt-out revolution for white college educated married women. In particular, the overall child penalty falls substantially between 1980 and 1990, and continues to fall between 1990 and 2000 although to a lesser extent, irrespective of the age group considered.<sup>29</sup>

Table 5 presents the decomposition of the overall child penalty by year and occupation. I focus on managerial/professional occupations for a number of reasons. First, managerial/professional occupations tend to be the focus of the media (see Williams et al., 2006). Second, lack of work place flexibility in these occupations (see Stone 2008) may result in women choosing to exit the labor market entirely on the one hand while on the other hand extreme hours in these types of occupations may afford women the ability to reduce their hours in light of new family responsibilities. This further highlights the importance of accounting for both the extensive and intensive when examining managerial/professional occupations.

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<sup>28</sup> Appendix Table 3 reveals that the results are very similar if I instead employ a cohort analysis, that is, I do not find evidence of the opt-out phenomenon among white college educated married women. Specifically, support of the opt-out phenomenon requires that the overall child penalty rises for more recent birth cohorts relative to earlier birth cohorts. I find, however, that the overall child penalty becomes smaller as one moves from earlier birth cohorts to more recent birth cohorts for all age groups for white married college educated married women.

<sup>29</sup> The patterns by age groups mirror those for the overall college educated sample for blacks, Hispanics, and Asians presented in Table 2. Results are available upon request.

Table 5 reveals that white college educated married women in managerial/professional occupations are more likely to adjust their hours (conditional on working) than exit the labor market, suggesting ignoring the intensive margin (as Belkin 2003 suggests) misses a large portion of the story. Despite this, it is still the case that white college educated married women in managerial/professional occupations experience a fall in the overall child penalty between 1990 and 2000 (36.29 to 32.05).

Looking at all managerial/professional occupations, however, combines all occupations irrespective of their gender mix. Given work-place flexibility and extreme hours likely vary a great deal by an occupation's gender mix, I look at detailed occupations in three occupation categories: male dominated occupations (architects, electrical engineers, physicians, lawyers, clergy/religious workers), female dominated occupations (registered nurses, dietitians/nutritionists, occupational therapists, primary school teachers, and librarians), and mixed occupations (human resources, accountants and auditors, secondary school teachers, economists, and editors/reporters). See the data section for a detailed discussion of occupation definitions.

The results are striking. For male dominated occupations the child penalty in unconditional annual hours generally falls between 1980 and 1990 and generally increases between 1990 and 2000.<sup>30</sup> For example, the child penalty in unconditional annual hours

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<sup>30</sup> The main exception is clergy/religious workers. This occupation category, as defined by the Census (<http://www.census.gov/hhes/www/ioindex/ioindex02/txtnew02.html#21-2099>) includes clergy ("Conduct religious worship and perform other spiritual functions associated with beliefs and practices of religious faith or denomination. Provide spiritual and moral guidance and assistance to members."), directors, religious activities and education ("Direct and coordinate activities of a denominational group to meet religious needs of students. Plan, direct, or coordinate church school programs designed to promote religious education among church membership. May provide counseling and guidance relative to marital, health, financial, and religious problems."), and all other religious workers ("All religious workers not listed separately."). The latter two categories are likely dominated by women, make up the bulk of the women in the overall clergy/religious workers occupation, and more closely resemble occupations in female dominated occupations (e.g., teachers). The overall results for clergy/religious workers, therefore, more closely align with female dominated, rather than male dominated occupations.



increased from 35.2 to 39.5 for architects between 1990 and 2000. This rise between 1990 and 2000 is in sharp contrast to the patterns found for female dominated occupations (with the exception of dietitians/nutritionists) and mixed occupations.<sup>31</sup> Perhaps not surprisingly, women are generally more likely to adjust their hours of work (conditional on working) in male dominated occupations. In female dominated occupations, the importance of the intensive margin varies depending on specialty. In the medical specialty (i.e., registered nurses, dietitians/nutritionists, occupational therapists) the intensive margin is of more importance than in the education specialty (i.e., primary school teachers, and librarians). For mixed occupations, the relative role of the extensive and intensive margin is closer together, particularly in earlier years. These patterns illustrate that accounting for differences in work-place flexibility and the incidence of extreme hours across occupations is important. Moreover, these patterns provide support for the popular press' notion of the opt-out revolution, particularly among white college educated married women in male dominated occupations.

## **6. Conclusion**

The “opt-out revolution” was first sensationalized in a 2003 *New York Times* article by Lisa Belkin and suggests that highly educated women (relative to their less educated counterparts) are exiting the labor market to care for their children more in recent years, that is, highly educated women face a higher child penalty than their less educated counterparts. This spurred a flurry of media attention to the issue, as well as increased interest among academics across the social sciences. The majority of the studies examine the employment side of the opt-

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<sup>31</sup> Analysis on alternative female dominated occupations and mixed occupations yield similar results and are available upon request.

out revolution (i.e., conditional on having children women exit the labor market).<sup>32</sup> With the exception of Vere (2007), they find little support for the media's account of highly educated women's recent exodus from the labor market to care for their children (see for example, Boushey 2005, Goldin 2006, Cohany and Sok 2007, Fortin 2008, and Percheski 2008).

The purpose of this study is to further document the evidence on the employment side of the opt-out revolution by allowing for differential effects of children on female labor supply by race, marital status, and occupation, as well as examining both the extensive margin (i.e., the decision to work a positive number of hours) and the intensive margin (i.e., conditional on participating in the labor market, the decision to adjust annual hours of work) of female labor supply.

Using the 1980, 1990, and 2000 U.S. Census, I find little evidence to support the opt-out revolution if one focuses solely on the extensive margin. In other words, I do not find that the child penalty for college educated women increased in recent years (1990 to 2000) irrespective of marital status or race. What is true, however, is that the child penalty falls substantially between 1980 and 1990 for all education levels and the penalty continues to fall (although to a lesser extent) for less educated women between 1990 and 2000 while the penalty remains flat during this time period for their college educated counterparts.

The patterns are very different if one focuses on the overall child penalty (based on both the extensive and intensive margins of labor supply). While the overall child penalty falls between 1990 and 2000 for white and Hispanic college educated married women, the opposite

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<sup>32</sup> There are several other aspects of the opt-out literature. Vere (2007), Percheski (2008), and Shang and Weinberg (2009) examine the fertility side of opting out. Goldin and Katz (2008), Herr and Wolfram (2009), and Bertrand, Goldin, and Katz, (2009) examine the impact of opting out on labor market outcomes among alumni at Harvard University and the University of Chicago's Booth School of Business, respectively. Stone and Lovejoy (2004), Hewlett *et al.* (2005), Williams *et al.* (2006), Hewlett (2007), and Stone (2008) seek to understand why highly educated women are exiting the labor market and what the associated costs of exiting may be.

pattern is found for their black and Asian counterparts. In addition, I find that the overall child penalty increases between 1990 and 2000 for white college educated married women who are in male dominated occupations, while the same is not true for female dominated occupations or mixed occupations. Overall, these patterns provide some evidence in support of the opt-out revolution, particularly for the group of women that tend to be the focus of the media, that is, white married college educated women in male dominated occupations.

Future work is needed to provide explanations for the patterns documented in this paper. Why do the opt-out patterns vary by race? Is it an artifact of cultural factors (i.e., different preferences for work and family)? Why do women opt-out? Is it family pulls (e.g., birth of a child), or is it work-place pushes (e.g., lack of flexibility on the job)? Or is it a combination of the two?<sup>33</sup> Does the importance of pulls vs. pushes vary by race? Conditional on exiting the labor market, how long do women remain out of the labor market? What types of occupations do they transition back into? How long do the transitions back into the labor market typically take?<sup>34</sup> Does race play a role in re-entry? Why do the opt-out patterns vary by occupation? Is it an artifact of differences in work- place flexibility and the incidence of extreme hours?

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<sup>33</sup> There is a growing literature that is investigating this, although it is generally focused on highly educated professional women (see Stone and Lovejoy 2004, Hewlett *et al.* 2005, Hewlett 2007, Stone 2008, and Herr and Wolfram 2009).

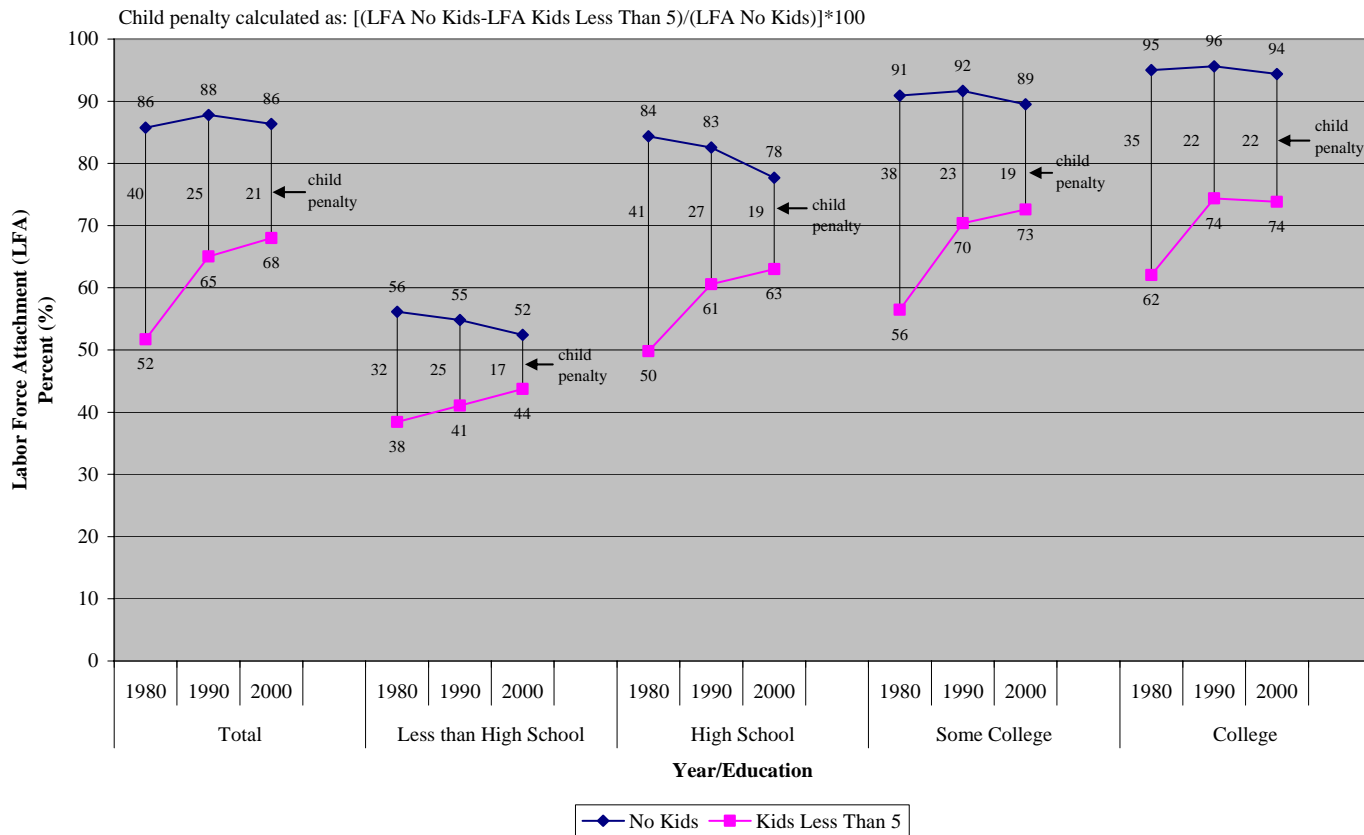
<sup>34</sup> Hewlett *et al.* (2005) and Hewlett (2007) provide evidence on re-entry for highly qualified professionals.

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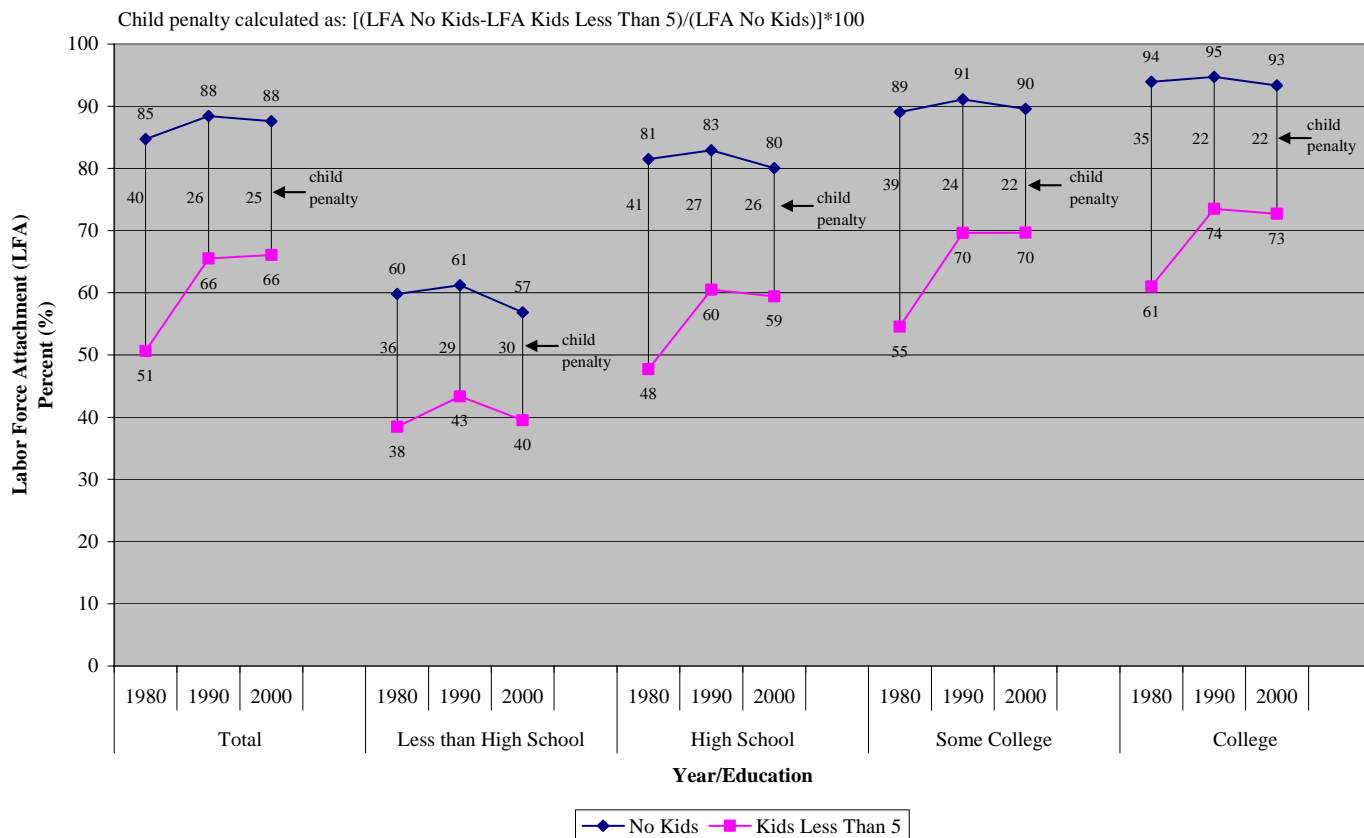
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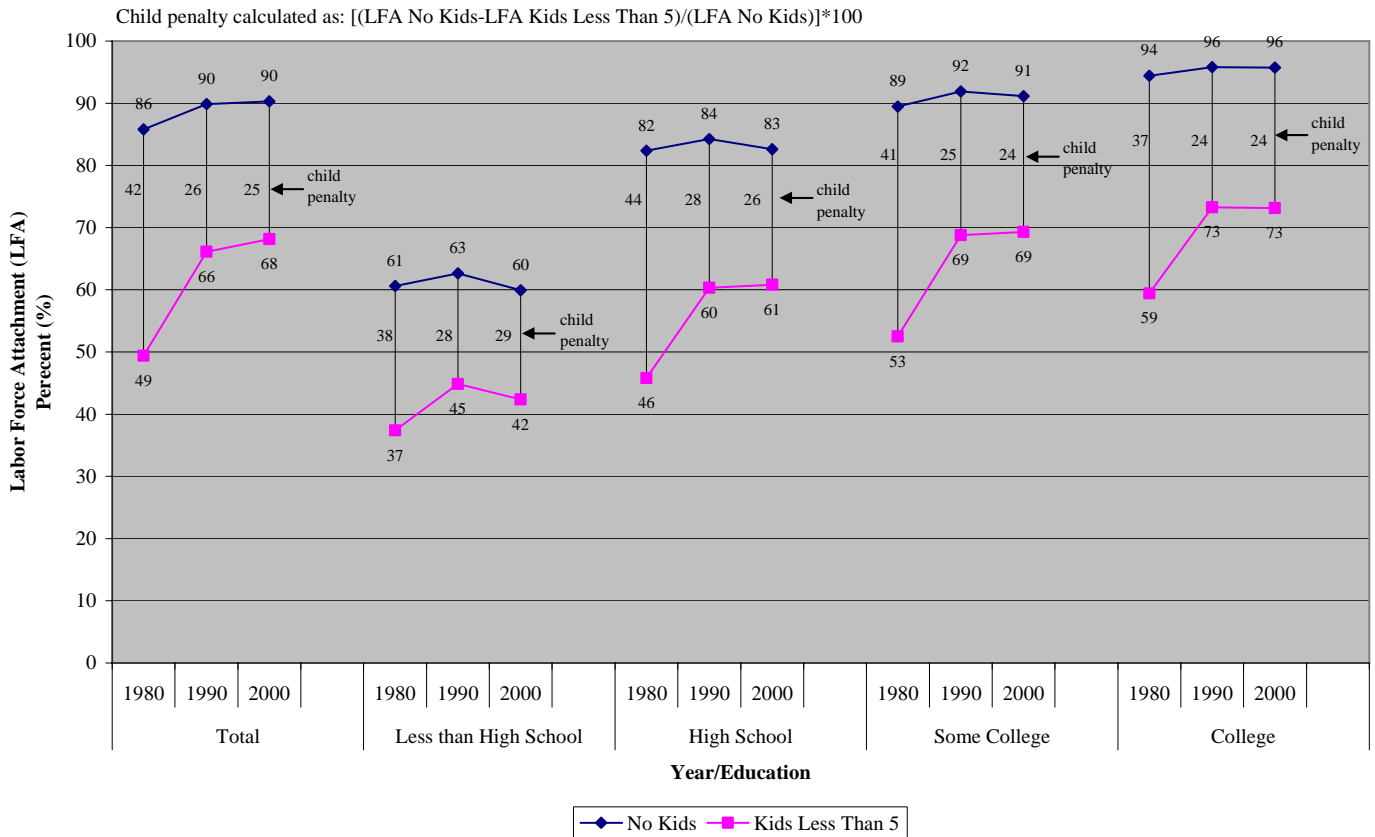
**Figure 1. Labor Force Attachment (LFA) by Presence of Young Children, Education, and Year  
(Women Aged 25-44)**



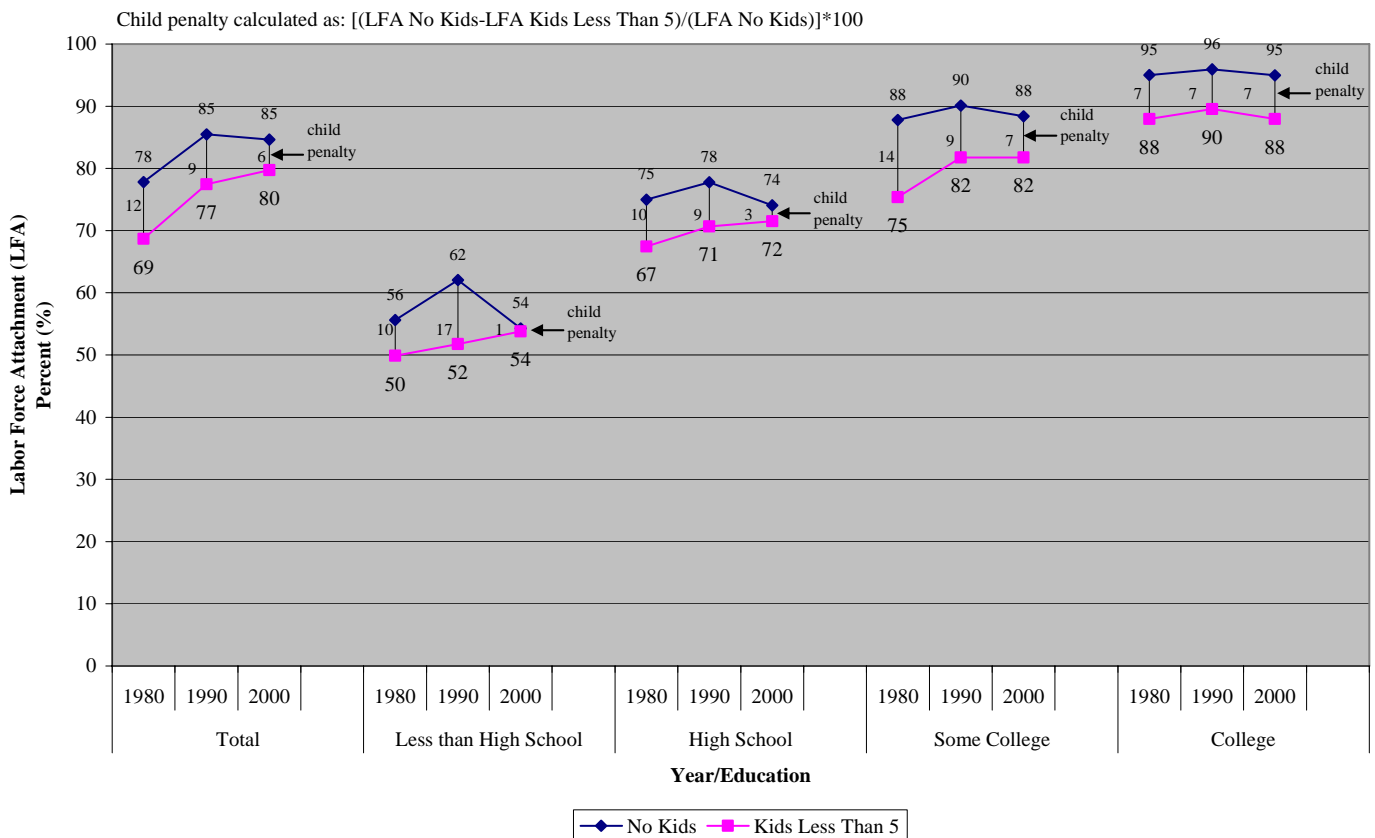
**Figure 2. Labor Force Attachment (LFA) by Presence of Children, Education, Year  
(Married Women Aged 25-44)**



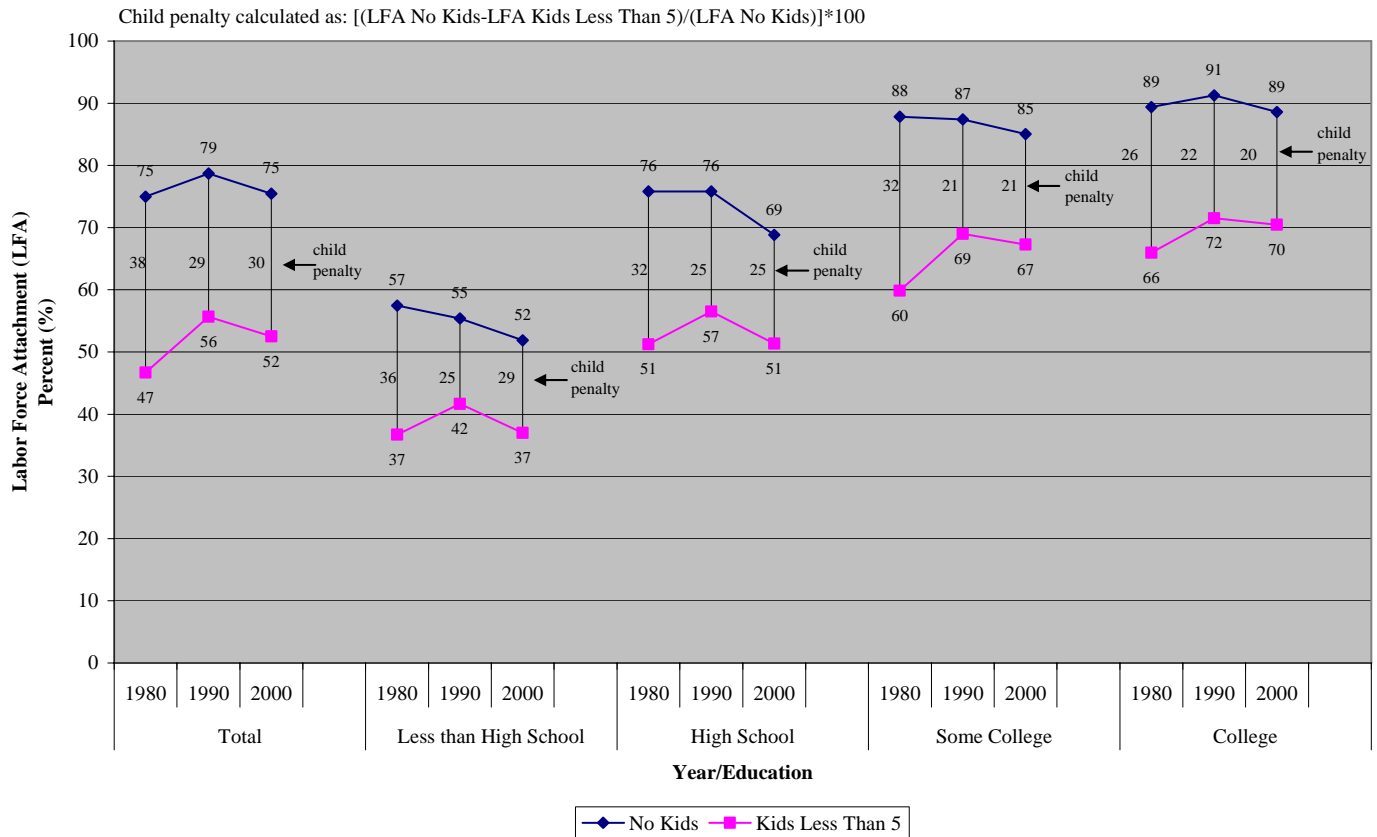
**Figure 3. Labor Force Attachment (LFA) by Presence of Young Children, Education, Year**  
(Married Non-Hispanic White Women Aged 25-44)



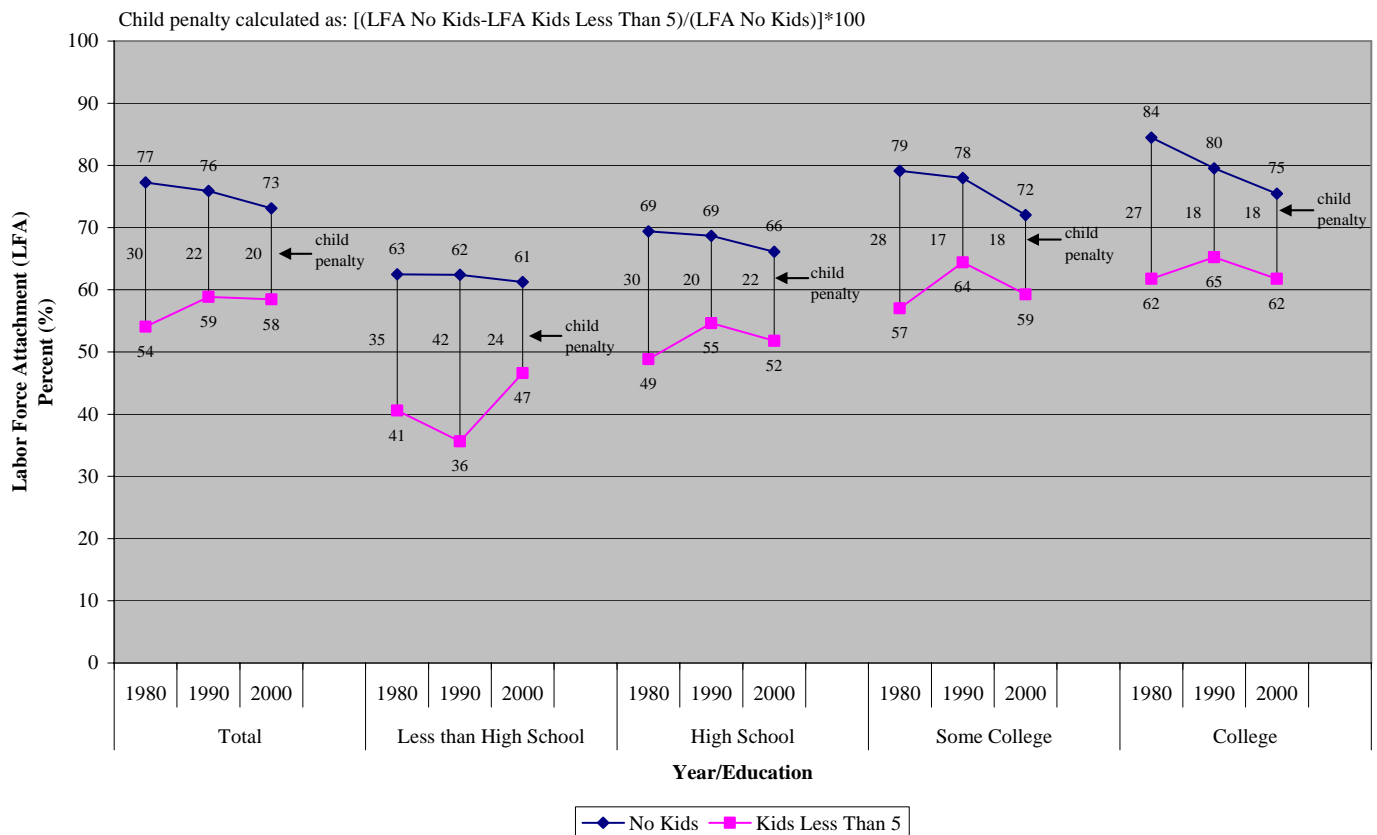
**Figure 4. Labor Force Attachment (LFA) by Presence of Young Children, Education, Year**  
(Married Non-Hispanic Black Women Aged 25-44)



**Figure 5. Labor Force Attachment (LFA) by Presence of Children, Education, Year  
(Married Hispanic Women Aged 25-44)**

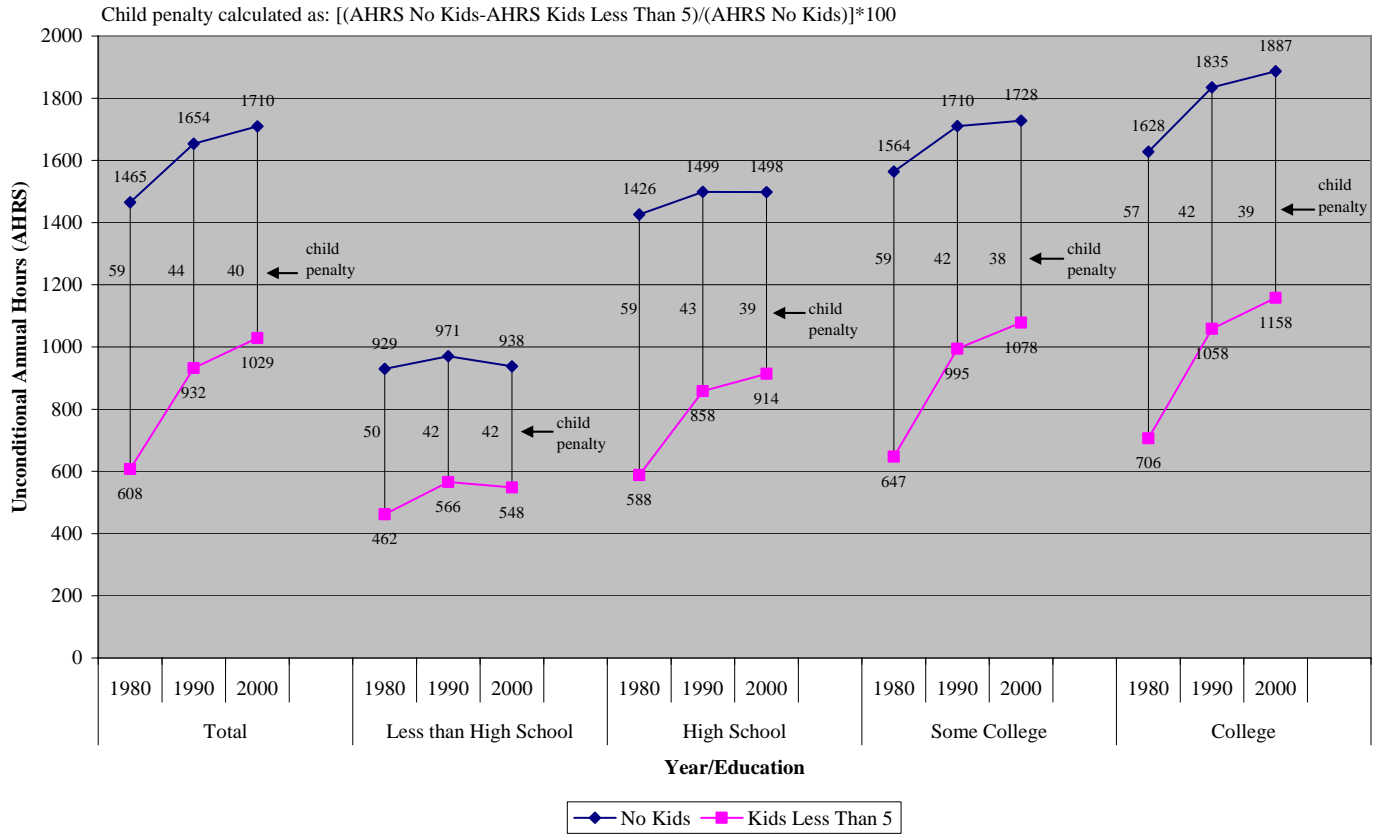


**Figure 6. Labor Force Attachment (LFA) by Presence of Young Children, Education, Year  
(Married Asian Women Aged 25-44)**

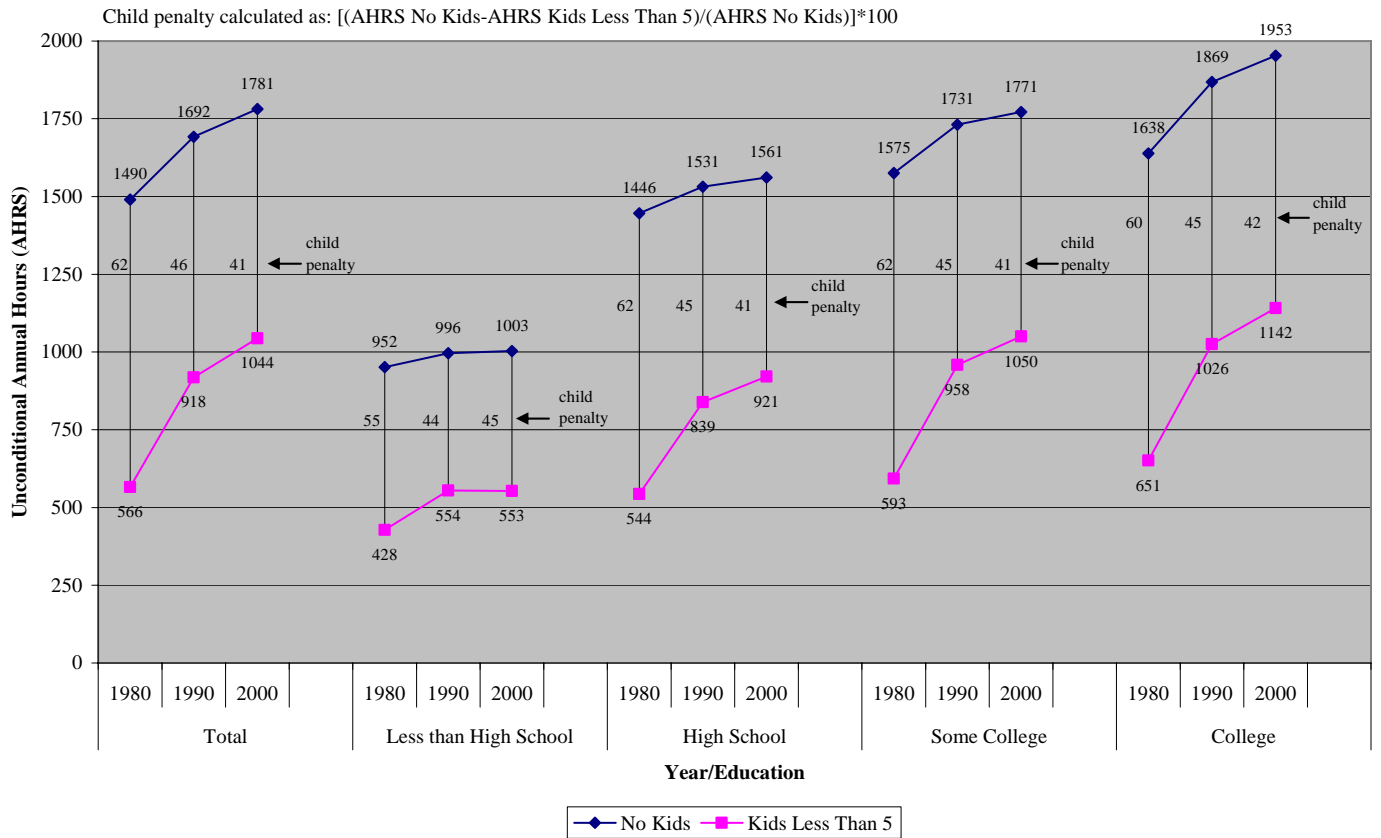




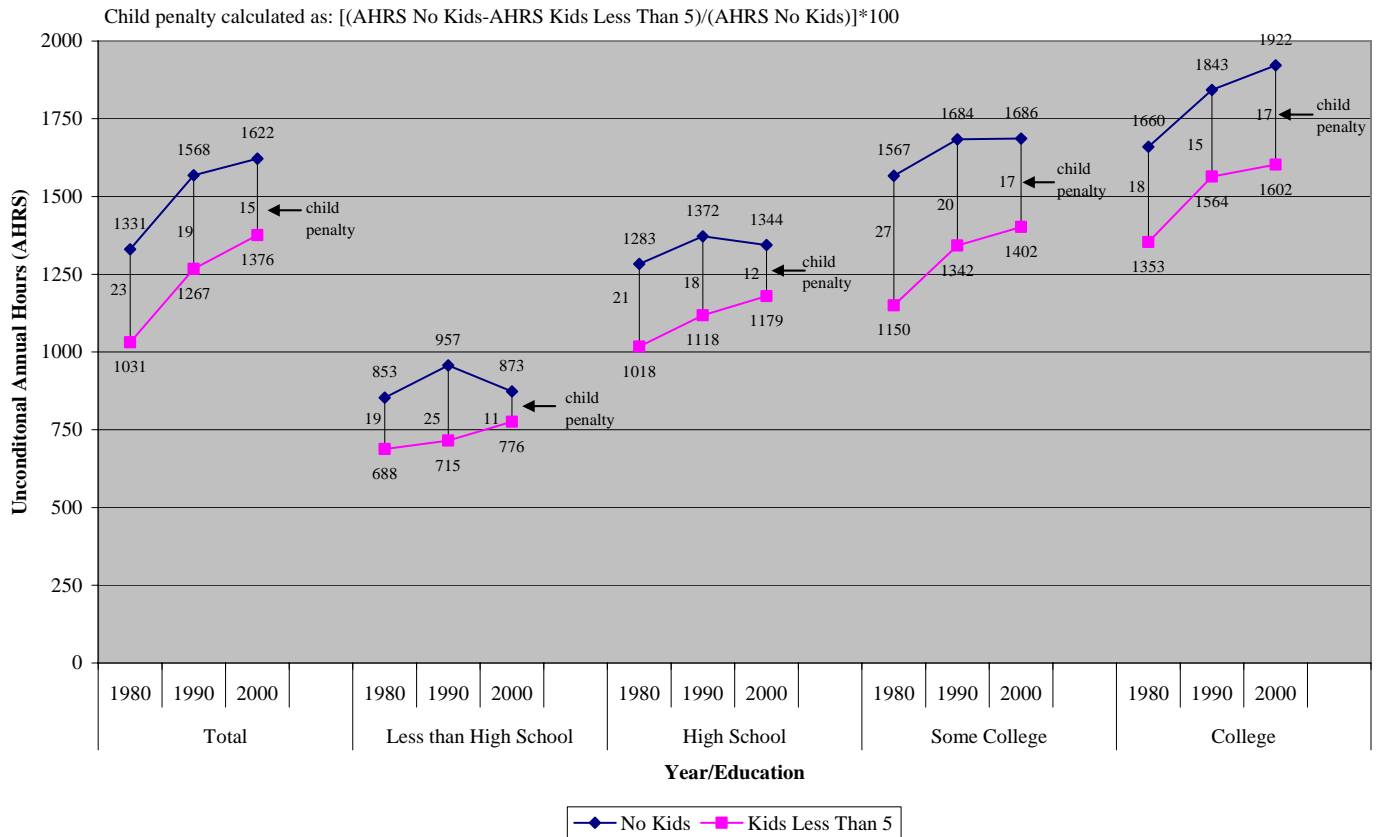
**Figure 7. Unconditional Annual Hours (AHRs) by Presence of Children, Education, Year (Married Women Aged 25-44)**



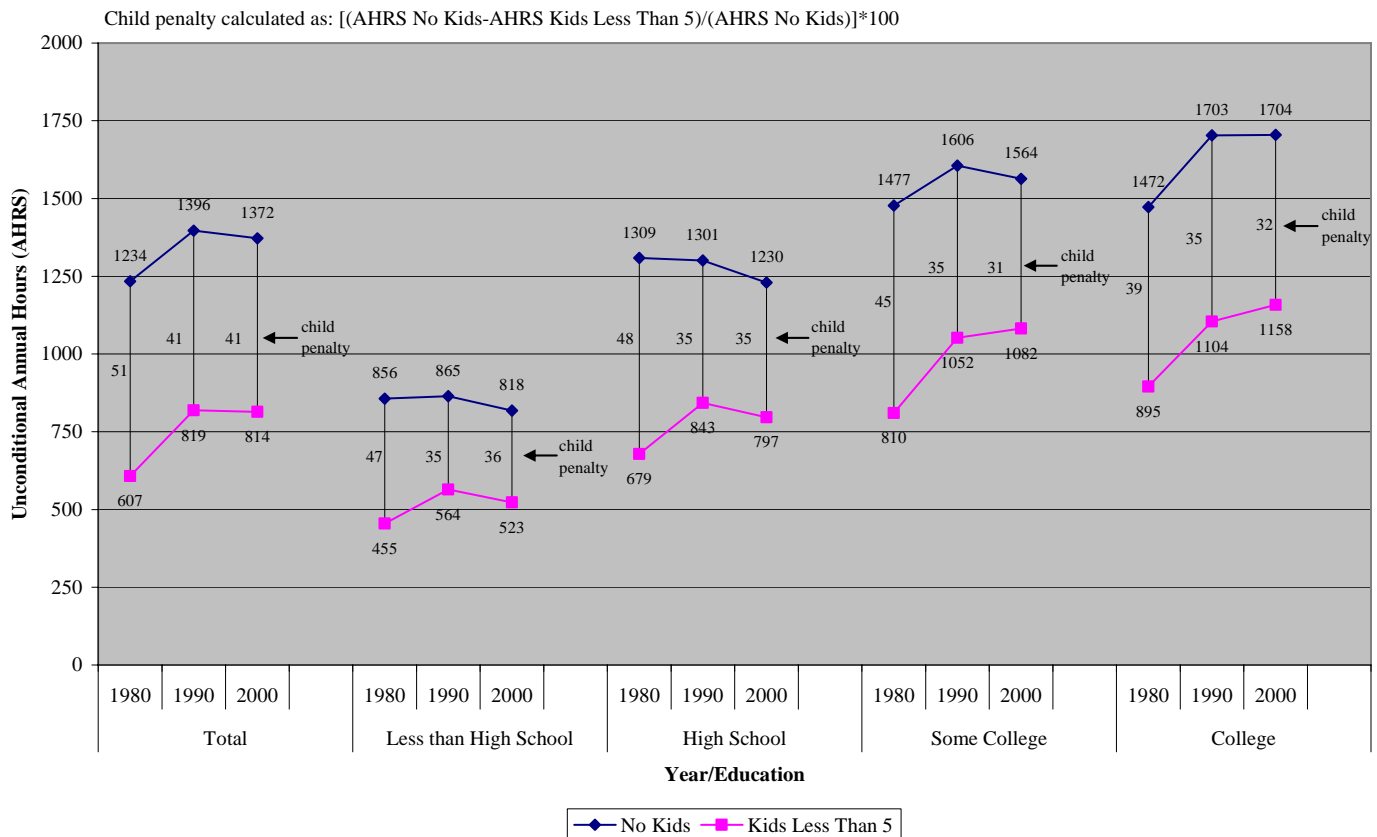
**Figure 8. Unconditional Annual Hours (AHRs) by Presence of Young Children, Education, Year (Married Non-Hispanic White Women Aged 25-44)**



**Figure 9. Unconditional Annual Hours (AHRs) by Presence of Young Children, Education, Year  
(Married Non-Hispanic Black Women Aged 25-44)**



**Figure 10. Unconditional Annual Hours (AHRs) by Presence of Children, Education, Year  
(Married Hispanic Women Aged 25-44)**



**Figure 11. Unconditional Annual Hours (AHRs) by Presence of Young Children, Education, Year  
(Married Asian Women Aged 25-44)**

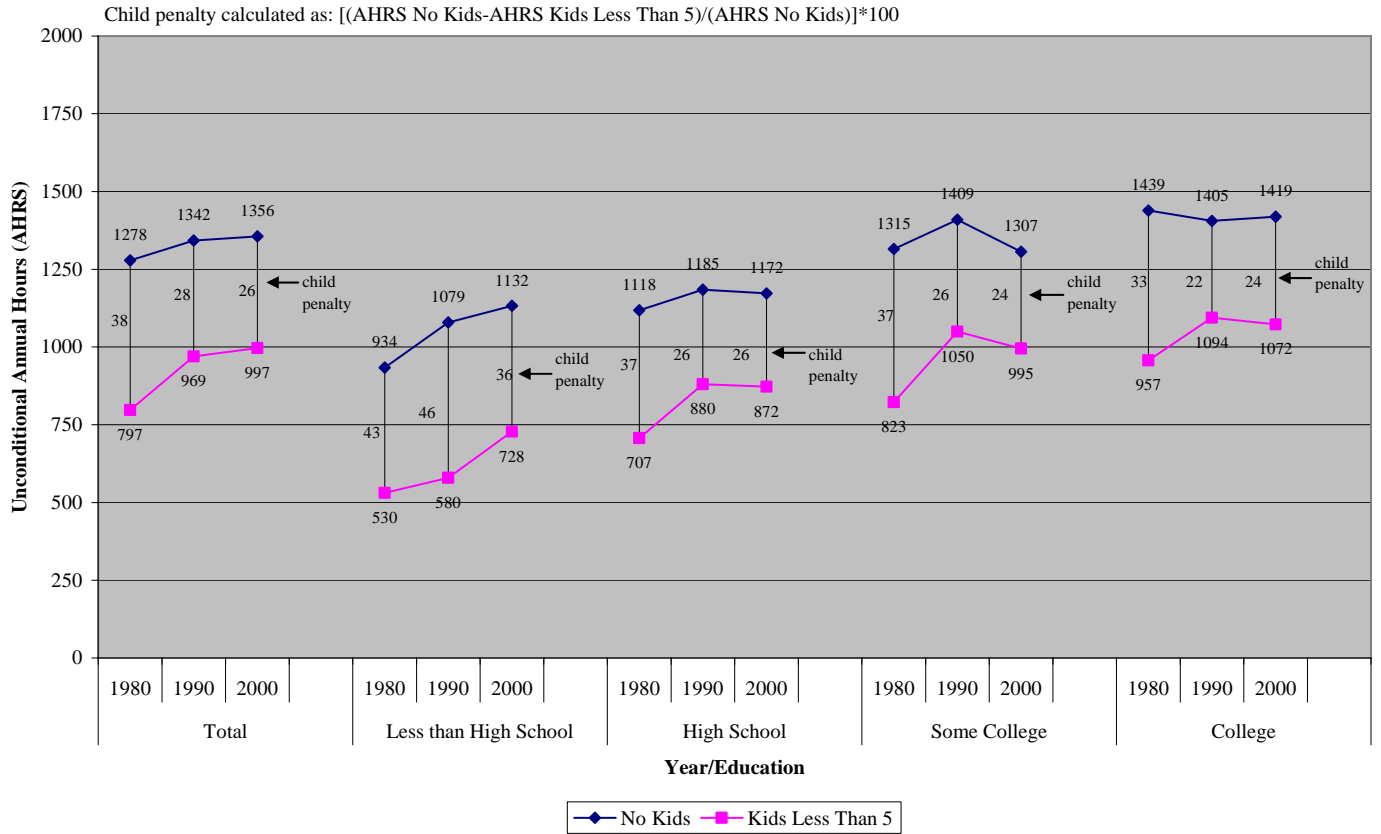


Table 1. Summary Statistics by Race/Ethnicity and Year  
(Women Aged 25-44)

	Total			Non-Hispanic White			Non-Hispanic Black			Hispanic			Asian		
	1980 (1)	1990 (2)	2000 (3)	1980 (4)	1990 (5)	2000 (6)	1980 (7)	1990 (8)	2000 (9)	1980 (10)	1990 (11)	2000 (12)	1980 (13)	1990 (14)	2000 (15)
<b>Race/Ethnicity<sup>^</sup></b>															
WHITE	0.81	0.77	0.71												
BLACK	0.10	0.11	0.11												
HISPANIC	0.06	0.08	0.11												
ASIAN	0.02	0.03	0.05												
<b>Labor Supply</b>															
AHRS	1032.74	1296.97	1393.57	1031.52	1327.86	1438.58	1131.59	1286.48	1444.46	866.78	1062.63	1125.51	1097.25	1244.07	1271.62
WKSWRK	28.10	33.91	35.76	28.24	34.93	37.11	29.96	32.82	36.30	23.16	27.49	28.92	28.47	31.23	31.63
UHRS	24.21	28.79	30.04	24.17	29.29	30.70	26.03	28.81	31.54	21.21	24.78	25.48	25.54	27.56	27.77
CAHRS*	1519.44	1679.63	1770.24	1507.02	1676.94	1775.69	1609.80	1718.35	1795.67	1495.89	1626.96	1689.56	1620.49	1760.41	1804.16
LFA	0.68	0.77	0.79	0.68	0.79	0.81	0.70	0.75	0.80	0.58	0.65	0.67	0.68	0.71	0.70
<b>Education</b>															
LTHIGH	0.18	0.11	0.09	0.15	0.07	0.05	0.29	0.15	0.10	0.46	0.35	0.32	0.18	0.13	0.08
HIGH	0.44	0.33	0.28	0.46	0.34	0.27	0.41	0.37	0.33	0.33	0.30	0.29	0.26	0.22	0.17
SOMECOL	0.19	0.32	0.34	0.20	0.33	0.35	0.19	0.33	0.38	0.13	0.24	0.25	0.19	0.25	0.23
COLLEGE	0.19	0.24	0.30	0.20	0.26	0.33	0.11	0.15	0.19	0.07	0.10	0.14	0.38	0.40	0.52
<b>Family Relationship</b>															
MARRIED	0.73	0.66	0.63	0.77	0.70	0.68	0.47	0.37	0.35	0.69	0.62	0.63	0.78	0.73	0.71
NCHILD	1.68	1.40	1.39	1.62	1.34	1.33	1.94	1.54	1.47	2.10	1.77	1.73	1.50	1.34	1.19
NCHLT5	0.35	0.35	0.34	0.34	0.34	0.33	0.34	0.32	0.29	0.48	0.44	0.42	0.44	0.38	0.36
NOKID	0.25	0.32	0.33	0.26	0.33	0.34	0.23	0.28	0.31	0.20	0.26	0.27	0.30	0.35	0.39
KIDLT5	0.28	0.27	0.27	0.27	0.27	0.26	0.27	0.25	0.23	0.37	0.34	0.33	0.35	0.30	0.29
<b>Age</b>															
AGE	33.40	34.11	34.95	33.45	34.23	35.21	33.18	33.79	34.70	33.17	33.37	33.90	33.30	34.21	34.33
N	1,296,068	1,635,258	1,589,181	1,052,561	1,283,288	1,152,947	131,166	158,876	164,045	76,955	126,889	165,967	26,514	51,220	65,403

Source: U.S. Census (IPUMS). All observations with flagged values on any of the variables of interest are dropped from the analysis. Sampling weights used.

See Appendix Table 1 for variable definitions.

<sup>^</sup> Does not sum to one as Native Americans and other races not included.

\*Number of observation is lower b/c they refer only to those working positive hours.

Table 2. Decomposition of the Child Penalty in Unconditional Annual Hours by Race/Ethnicity, Year, and Education  
(Married Women Aged 25-44)

	Total			Non-Hispanic White			Non-Hispanic Black			Hispanic			Asian		
	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)
<b>Total</b>															
Child Penalty	58.53	43.64	39.84	62.01	45.71	41.42	22.54	19.20	15.17	50.77	41.34	40.69	37.64	27.77	26.47
Extensive Margin	40.20	25.89	24.55	42.41	26.40	24.53	11.71	9.40	5.80	37.73	29.26	30.41	30.01	22.46	20.02
% Explained	(68.68)	(59.32)	(61.62)	(68.40)	(57.74)	(59.22)	(51.93)	(48.94)	(38.21)	(74.31)	(70.79)	(74.74)	(79.73)	(80.87)	(75.65)
Intensive Margin	18.34	17.76	15.29	19.60	19.32	16.89	10.84	9.80	9.37	13.04	12.07	10.28	7.63	5.31	6.44
% Explained	(31.32)	(40.68)	(38.38)	(31.60)	(42.26)	(40.78)	(48.07)	(51.06)	(61.79)	(25.69)	(29.21)	(25.26)	(20.27)	(19.13)	(24.35)
<b>Less Than High School</b>															
Child Penalty	50.29	41.69	41.54	55.01	44.36	44.89	19.39	25.26	11.14	46.90	34.73	36.06	43.19	46.30	35.72
Extensive Margin	35.65	29.22	30.48	38.30	28.37	29.28	10.30	16.61	0.89	36.07	24.78	28.69	35.07	42.90	23.91
% Explained	(70.89)	(70.08)	(73.39)	(69.63)	(63.96)	(65.23)	(53.13)	(65.77)	(7.97)	(76.91)	(71.35)	(79.57)	(81.21)	(92.66)	(66.93)
Intensive Margin	14.64	12.47	11.05	16.71	15.99	15.61	9.09	8.65	10.26	10.83	9.95	7.37	8.12	3.40	11.81
% Explained	(29.11)	(29.92)	(26.61)	(30.37)	(36.04)	(34.77)	(46.87)	(34.23)	(92.03)	(23.09)	(28.65)	(20.43)	(18.79)	(7.34)	(33.07)
<b>High School</b>															
Child Penalty	58.76	42.74	39.00	62.39	45.24	40.97	20.70	18.48	12.24	48.16	35.18	35.22	36.76	25.70	25.58
Extensive Margin	41.43	27.05	25.79	44.38	28.38	26.41	10.10	9.15	3.45	32.42	25.49	25.38	29.58	20.43	21.63
% Explained	(70.51)	(63.29)	(66.13)	(71.13)	(62.74)	(64.46)	(48.80)	(49.50)	(28.15)	(67.33)	(72.45)	(72.06)	(80.47)	(79.52)	(84.58)
Intensive Margin	17.33	15.69	13.21	18.01	16.86	14.56	10.60	9.33	8.80	15.73	9.69	9.84	7.18	5.26	3.94
% Explained	(29.49)	(36.71)	(33.87)	(28.87)	(37.26)	(35.54)	(51.20)	(50.50)	(71.85)	(32.67)	(27.55)	(27.94)	(19.53)	(20.48)	(15.42)
<b>Some College</b>															
Child Penalty	58.65	41.84	37.61	62.36	44.65	40.71	26.58	20.28	16.83	45.15	34.50	30.80	37.44	25.53	23.85
Extensive Margin	38.75	23.58	22.22	41.29	25.13	23.97	14.14	9.28	7.48	31.84	21.06	20.88	27.93	17.42	17.73
% Explained	(66.08)	(56.36)	(59.09)	(66.21)	(56.29)	(58.88)	(53.18)	(45.77)	(44.41)	(70.52)	(61.04)	(67.78)	(74.61)	(68.24)	(74.35)
Intensive Margin	19.90	18.26	15.39	21.07	19.52	16.74	12.45	11.00	9.36	13.31	13.44	9.92	9.51	8.11	6.12
% Explained	(33.92)	(43.64)	(40.91)	(33.79)	(43.71)	(41.12)	(46.82)	(54.23)	(55.59)	(29.48)	(38.96)	(32.22)	(25.39)	(31.76)	(25.65)
<b>College</b>															
Child Penalty	56.61	42.31	38.62	60.25	45.11	41.55	18.49	15.13	16.63	39.20	35.16	32.04	33.47	22.11	24.42
Extensive Margin	35.05	22.40	22.11	37.06	23.53	23.60	7.45	6.65	7.39	26.18	21.65	20.44	26.85	17.94	18.11
% Explained	(61.92)	(52.94)	(57.24)	(61.52)	(52.16)	(56.78)	(40.29)	(43.96)	(44.42)	(66.79)	(61.59)	(63.80)	(80.22)	(81.16)	(74.15)
Intensive Margin	21.56	19.91	16.51	23.19	21.58	17.96	11.04	8.48	9.24	13.02	13.50	11.60	6.62	4.17	6.31
% Explained	(38.08)	(47.06)	(42.76)	(38.48)	(47.84)	(43.22)	(59.71)	(56.04)	(55.58)	(33.21)	(38.41)	(36.20)	(19.78)	(18.84)	(25.85)

Source: U.S. Census (IPUMS). All observations with flagged values on any of the variables of interest are dropped from the analysis. Sampling weights used. Numbers do not always add up due to rounding.

Table 3. Decomposition of the Child Penalty in Unconditional Annual Hours by Year  
(Non-Hispanic White College Educated Married Women)

	1980 (1)	1981-1983 (2)	1984-1986 (3)	1987-1989 (4)	1990 (5)	1991-1993 (6)	1994-1996 (7)	1997-1999 (8)	2000 (9)	2001-2003 (10)	2004-2006 (11)	2007-2009 (12)
Child Penalty	60.25	55.09	48.23	46.66	45.11	41.44	40.03	40.17	41.55	42.22	39.90	38.31
Extensive Margin	37.06	30.07	25.95	24.56	23.53	20.66	19.98	20.13	23.60	23.46	22.56	21.09
% Explained	(61.52)	(54.59)	(53.81)	(52.65)	(52.16)	(49.84)	(49.90)	(50.12)	(56.78)	(55.57)	(56.54)	(55.04)
Intensive Margin	23.19	25.02	22.27	22.09	21.58	20.78	20.05	20.04	17.96	18.75	17.34	17.22
% Explained	(38.48)	(45.41)	(46.19)	(47.35)	(47.84)	(50.16)	(50.10)	(49.88)	(43.22)	(44.43)	(43.46)	(44.96)

Source: U.S. Census (IPUMS) in yellow and the March Current Population Survey (IPUMS). All observations with flagged values on any of the variables of interest are dropped from the analysis. Sampling weights used. Numbers do not always add up due to rounding.

Table 4. Decomposition of the Child Penalty in Unconditional Annual Hours by Year and Age  
(Non-Hispanic White College Educated Married Women)

	Total			25-29			30-34			35-39			40-44		
	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)	(15)
<b>Total</b>															
Child Penalty	60.25	45.11	41.55	54.46	38.85	34.12	62.47	46.14	42.77	64.20	48.33	45.47	59.95	44.29	41.89
Extensive Margin	37.06	23.53	23.60	29.50	16.97	16.01	39.23	23.73	23.25	43.46	26.50	26.96	41.00	25.85	24.90
% Explained	(61.52)	(52.16)	(56.78)	(54.16)	(43.68)	(46.91)	(62.81)	(51.43)	(54.35)	(67.70)	(54.84)	(59.29)	(68.39)	(58.37)	(59.44)
Intensive Margin	23.19	21.58	17.96	24.96	21.88	18.12	23.23	22.41	19.53	20.74	21.83	18.51	18.95	18.44	16.99
% Explained	(38.48)	(47.84)	(43.22)	(45.84)	(56.32)	(53.09)	(37.19)	(48.57)	(45.65)	(32.30)	(45.16)	(40.71)	(31.61)	(41.63)	(40.56)

Source: U.S. Census (IPUMS). All observations with flagged values on any of the variables of interest are dropped from the analysis. Sampling weights used. Numbers do not always add up due to rounding.

Table 5. Decomposition of the Child Penalty in Unconditional Annual Hours by Year and Occupation  
(Non-Hispanic White College Educated Married Women)

	Managerial/Professional														
	1980	1990	2000												
Child Penalty	49.19	36.29	32.05												
Extensive Margin	21.89	13.02	12.39												
% Explained	(44.50)	(35.89)	(38.67)												
Intensive Margin	27.30	23.27	19.65												
% Explained	(55.50)	(64.11)	(61.33)												
<b>Traditionally Male Dominated Occupations</b>															
	Architects			Electrical Engineers			Physicians			Lawyers			Clergy/Religious Workers		
	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000
Child Penalty	51.44	35.16	39.51	27.25	25.66	27.20	28.46	19.88	23.81	24.02	24.53	26.97	46.99	47.77	43.24
Extensive Margin	10.60	9.35	15.12	25.12	9.05	14.11	6.27	3.24	2.22	8.10	8.89	7.97	18.94	14.33	9.66
% Explained	(20.61)	(26.58)	(38.26)	(92.20)	(35.25)	(51.85)	(22.03)	(16.28)	(9.32)	(33.73)	(36.23)	(29.57)	(40.30)	(29.99)	(22.34)
Intensive Margin	40.84	25.82	24.39	2.13	16.62	13.10	22.19	16.64	21.59	15.92	15.64	18.99	28.05	33.45	33.58
% Explained	(79.39)	(73.42)	(61.74)	(7.80)	(64.75)	(48.15)	(77.97)	(83.72)	(90.68)	(66.27)	(63.77)	(70.43)	(59.70)	(70.01)	(77.66)
<b>Traditionally Female Dominated Occupations</b>															
	Registered Nurses			Dietitians/Nutritionists			Occupational Therapists			Primary School Teachers			Librarians		
	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000
Child Penalty	49.97	38.05	32.99	48.69	38.44	40.63	49.75	38.38	33.53	48.33	32.26	28.94	49.48	36.48	34.67
Extensive Margin	16.40	9.30	9.56	19.92	9.56	13.43	20.40	6.68	8.04	25.65	14.08	14.16	25.93	16.59	14.78
% Explained	(32.81)	(24.43)	(28.97)	(40.90)	(24.86)	(33.06)	(41.01)	(17.40)	(23.99)	(53.08)	(43.64)	(48.92)	(52.40)	(45.46)	(42.63)
Intensive Margin	33.57	28.76	23.44	28.78	28.89	27.20	29.35	31.70	25.48	22.68	18.18	14.78	23.55	19.89	19.89
% Explained	(67.19)	(75.57)	(71.03)	(59.10)	(75.14)	(66.94)	(58.99)	(82.60)	(76.01)	(46.92)	(56.36)	(51.08)	(47.60)	(54.54)	(57.37)
<b>Mixed Occupations</b>															
	Human Resources			Accountants/Auditors			Secondary School Teachers			Economists			Editors/Reporters		
	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000	1980	1990	2000
Child Penalty	38.57	30.46	26.37	43.03	35.09	33.31	45.49	31.72	29.61	41.18	38.93	28.85	56.35	42.27	40.89
Extensive Margin	29.02	17.29	14.35	23.05	15.04	13.00	21.67	13.82	11.80	23.30	21.31	10.81	23.28	13.70	12.69
% Explained	(75.22)	(56.77)	(54.42)	(53.55)	(42.86)	(39.02)	(47.65)	(43.56)	(39.83)	(56.58)	(54.75)	(37.45)	(41.31)	(32.42)	(31.04)
Intensive Margin	9.56	13.17	12.02	19.99	20.05	20.31	23.81	17.90	17.82	17.88	17.62	18.05	33.07	28.57	28.20
% Explained	(24.78)	(43.23)	(45.58)	(46.45)	(57.14)	(60.98)	(52.35)	(56.44)	(60.17)	(43.42)	(45.25)	(62.55)	(58.69)	(67.58)	(68.96)

Source: U.S. Census (IPUMS). All observations with flagged values on any of the variables of interest are dropped from the analysis. Sampling weights used. Numbers do not always add up due to rounding.

See Appendix Table 2 for percent male in the occupation and sample sizes.



Appendix Table 1. Variable Definitions

Variable	Definition
<b>Labor Supply</b>	
UAHRS	=unconditional annual hours=(WKSWRK*UHRS)
WKSWRK	=weeks worked last year
UHRS	=usual hours worked per week
CAHRS	=conditional annual hours=UAHRS if UAHRS>0
LFA	=labor force attachment=1 if UAHRS>0, 0 otherwise
LFP	=labor force participation=1 if EMP or UNEMP, 0 otherwise
EMP	=1 if employed, 0 otherwise
UNEMP	=1 if unemployed, 0 otherwise
<b>Education</b>	
LTHIGH	=1 if less than high school, 0 otherwise
HIGH	=1 if high school (with or without degree), 0 otherwise
SOMECOL	=1 if some college, 0 otherwise
COLLEGE	=1 if ba or higher, 0 otherwise
Race/Ethnicity	
WHITE	=1 if non-Hispanic white, 0 otherwise
BLACK	=1 if non-Hispanic black, 0 otherwise
HISPANIC	=1 if Hispanic, 0 otherwise
ASIAN	=1 if Asian, 0 otherwise
NATIVE	=1 if Native American, 0 otherwise
OTHRACE	=1 if other race, 0 otherwise
<b>Family Relationship</b>	
MARRIED	=1 if married spouse present, 0 otherwise
NCHILD	=number of own children in the household
NCHLT5	=number of own children less than 5 in the household
NOKID	=1 if no own children present in the household, 0 otherwise
KIDLT5	=1 if any own children less than 5 present in the household, 0 otherwise
<b>Age</b>	
AGE	=age

Appendix Table 2. Descriptive Statistics for Detailed Occupations by Occupation Type and Year

	1980			1990			2000		
	%Male	N (No Kids)	N (Kids lt5)	%Male	N (No Kids)	N (Kids lt5)	%Male	N (No Kids)	N (Kids lt5)
<b>Male Dominated Occupations</b>									
Chief executives and public administrat	0.747	14	8	0.703	4	5	0.800	381	487
Funeral directors	0.912	3	2	0.864	6	4	0.819	7	10
Purchasing agents and buyers, of farm p	0.904	0	2	0.799	2	2	0.754	4	2
Construction inspectors	0.948	1	0	0.929	3	7	0.890	3	5
Architects	0.905	64	25	0.835	155	169	0.778	178	183
Aerospace engineer	0.966	10	0	0.913	76	52	0.904	34	64
Metallurgical and materials engineers, Petroleum, mining, and geological engin	0.947	8	1	0.865	9	14	0.886	16	10
Petroleum, mining, and geological engin	0.954	2	1	0.934	11	11	0.928	7	11
Chemical engineers	0.946	12	6	0.882	50	32	0.845	52	64
Civil engineers	0.966	26	12	0.926	90	96	0.888	132	137
Electrical engineer	0.947	52	25	0.899	238	179	0.892	115	114
Industrial engineers	0.898	44	13	0.853	84	77	0.825	122	135
Mechanical engineers	0.976	14	4	0.945	48	51	0.932	75	91
Not-elsewhere-classified engineers	0.957	34	25	0.896	221	174	0.880	178	179
Physicists and astronomers	0.940	7	4	0.852	18	13	0.860	10	10
Atmospheric and space scientists	0.858	1	3	0.877	7	1	0.878	10	5
Geologists	0.871	35	10	0.834	79	53	0.742	133	111
Agricultural and food scientists	0.752	17	10	0.707	35	42	0.726	33	34
Foresters and conservation scientists	0.877	16	8	0.852	36	20	0.841	36	33
Physicians	0.849	257	207	0.778	562	853	0.715	629	1142
Dentists	0.918	30	23	0.857	108	170	0.802	95	159
Optometrists	0.902	9	3	0.821	33	44	0.693	52	77
Podiatrists	0.909	0	2	0.891	5	5	0.821	10	17
Clergy and religious workers	0.851	78	138	0.766	153	63	0.702	214	361
Lawyers	0.844	542	262	0.734	1310	1333	0.682	1350	1728
Announcers	0.790	16	7	0.780	24	29	0.764	16	14

Appendix Table 2. Descriptive Statistics for Detailed Occupations by Occupation Type and Year (Continued)

	1980			1990			2000		
	%Male	N (No Kids)	N (Kids lt5)	%Male	N (No Kids)	N (Kids lt5)	%Male	N (No Kids)	N (Kids lt5)
<b>Female Dominated Occupations</b>									
Registered nurses	0.039	2027	2853	0.054	3099	6284	0.074	2835	5784
Dietitians and nutritionists	0.096	183	180	0.099	225	417	0.094	211	350
Occupational therapists	0.073	106	94	0.113	206	329	0.096	313	495
Physical therapists	0.257	186	185	0.236	342	507	0.269	520	830
Speech therapists	0.099	329	386	0.072	382	710	0.069	512	722
Therapists, n.e.c.	0.299	106	86	0.259	214	239	0.238	264	270
Kindergarten and earlier school teacher	0.031	410	1058	0.019	254	1037	0.019	508	1257
Primary school teachers	0.219	8251	12576	0.204	6989	11640	0.200	6993	11536
Special education teachers	0.300	89	85	0.163	164	314	0.128	522	688
Librarians	0.167	533	401	0.196	382	403	0.176	256	242
Dancers	0.218	10	4	0.212	16	17	0.130	8	11
<b>Mixed Occupations</b>									
Legislators	0.614	5	11	0.573	14	10	0.555	12	8
Human resources and labor relations man	0.615	184	94	0.493	365	267	0.432	540	607
Managers in education and related field	0.596	422	308	0.466	726	813	0.384	675	844
Postmasters and mail superintendents	0.572	1	2	0.540	4	11	0.445	6	7
Managers of properties and real estate	0.550	43	56	0.507	197	209	0.454	181	235
Accountants and auditors	0.588	894	537	0.453	2768	2752	0.405	2688	3967
Other financial specialists	0.516	313	210	0.461	739	742	0.489	734	899
Personnel, HR, training, and labor rela	0.504	491	275	0.416	757	662	0.325	1419	1450
Buyers, wholesale and retail trade	0.517	158	71	0.439	207	197	0.434	185	224
Medical scientists	0.579	32	28	0.558	75	68	0.524	119	103
Subject instructors (HS/college)	0.590	878	1156	0.559	1023	1237	0.510	1460	1684
Secondary school teachers	0.405	2586	3380	0.414	1386	1907	0.396	1462	1975
Archivists and curators	0.504	41	12	0.440	69	44	0.411	69	62
Economists, market researchers, and sur	0.691	151	75	0.549	376	289	0.513	203	238
Social scientists, n.e.c.	0.576	26	16	0.518	44	41	0.485	72	51
Writers and authors	0.507	79	77	0.480	200	281	0.439	247	405
Designers	0.453	380	296	0.408	817	831	0.416	968	1078
Art makers: painters, sculptors, craft-	0.484	223	159	0.446	308	77	0.505	258	220
Editors and reporters	0.477	429	306	0.467	576	554	0.489	506	487

Source: U.S. Census (IPUMS). All observations with flagged values on any of the variables of interest are dropped from the analysis. Sampling weights used.

Notes: % Male is the percent of the occupation that is male for all individuals between the ages of 18-64. N (No Kids) and N(Kids lt5) are the sample sizes for college educated Non-Hispanic white married women between the ages of 25-44 without and with kids less than 5, respectively.

Appendix Table 3. Decomposition of the Child Penalty in Unconditional Annual Hours by Birth Cohort and Age

<b>Overall Child Penalty</b>				
	Age			
	25-29	30-34	35-39	40-44
<b>Birth Cohort</b>				
1936-1945			64.203	59.954
1946-1955	54.463	62.466	48.330	44.291
1956-1965	38.851	46.142	45.475	41.887
1966-1975	34.123	42.772		
<b>Extensive Margin</b>				
	Age			
	25-29	30-34	35-39	40-44
<b>Birth Cohort</b>				
1936-1945			43.463	41.002
1946-1955	29.499	39.233	26.502	25.853
1956-1965	16.972	23.730	26.960	24.897
1966-1975	16.005	23.246		
<b>Intensive Margin</b>				
	Age			
	25-29	30-34	35-39	40-44
<b>Birth Cohort</b>				
1936-1945			20.741	18.952
1946-1955	24.964	23.233	21.828	18.437
1956-1965	21.879	22.412	18.514	16.990
1966-1975	18.118	19.526		

Source: U.S. Census (IPUMS). All observations with flagged values on any of the variables of interest are dropped from the analysis. Sampling weights used. Numbers do not always add up due to rounding.