

Carstensen, Kai; Gundlach, Erich; Hartmann, Susanne

**Working Paper**

## The augmented Solow model with Mincerian schooling and externalities

Kiel Working Paper, No. 1408

**Provided in Cooperation with:**

Kiel Institute for the World Economy – Leibniz Center for Research on Global Economic Challenges

*Suggested Citation:* Carstensen, Kai; Gundlach, Erich; Hartmann, Susanne (2008) : The augmented Solow model with Mincerian schooling and externalities, Kiel Working Paper, No. 1408, Kiel Institute for the World Economy (IfW), Kiel

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## **The Augmented Solow Model with Mincerian Schooling and Externalities**

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**No. 1408** | March, 2008

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## The Augmented Solow Model with Mincerian Schooling and Externalities \*

Kai Carstensen, Erich Gundlach, and Susanne Hartmann

### Abstract:

We combine the augmented Solow model with the Mincer equation to derive a specification that identifies an education externality within a production function framework. The previous empirical literature has not reached a consensus about the size of the education externality, which is given by the difference between the microeconomic and the macroeconomic return to education. Relative to our benchmark value that is based on a parameterization of the derived specification, we find that the estimated education externality is too large when the empirical model is not properly restricted, and appears to be absent when all control variables of the empirical model are properly accounted for. We note that the absence of an education externality would be inconsistent with observed levels education subsidies.

Keywords: Augmented Solow model, education externalities, Mincer equation

JEL classification: O4

### Kai Carstensen

ifo Institut für Wirtschaftsforschung  
Postfach 86 04 60  
81631 München  
Phone: +49 89 9224-1266  
Fax: +49 89 985369  
Email: [carstensen@ifo.de](mailto:carstensen@ifo.de)

### Susanne Hartmann

Deutsche Gesellschaft für technische  
Zusammenarbeit (GTZ)  
P.O. Box 5180  
65726 Eschborn  
email: [susanne.hartmann@gtz.de](mailto:susanne.hartmann@gtz.de)

### Erich Gundlach

Kiel Institute for the World Economy  
24100 Kiel, Germany  
Telephone: +49 431 8814 284  
E-mail: [erich.gundlach@ifw-kiel.de](mailto:erich.gundlach@ifw-kiel.de)

\* We are grateful for comments on an earlier version to participants of a Kiel Institute-MPI Workshop, Werner Bönnte, two anonymous referees, and Bjarne S. Jensen.

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## **1. Introduction**

Educational capital appears to be one of the factors that are closely related to long-run economic development. For instance, a high level of education per worker may increase the productivity of other production factors like physical capital and co-workers, or a high aggregate level of educational capital may generate more ideas and hence faster technological change. Thus educational capital accumulation may have macroeconomic returns that go beyond the individual private returns. Presumed education externalities in turn appear to justify the prevailing public subsidization of education in most if not all countries of the world.

Positive education externalities should show up as the difference between the macroeconomic and the microeconomic return to education. Empirical work on this issue has not derived clear-cut conclusions up to now. Some authors find evidence in favor of aggregate education externalities, but others deny their empirical relevance, as summarized by Pritchett (2006). We think that part of the inconclusive state of art has to do with the occasional comparison of returns that is inconsistent with a simple production function framework, and with the neglect of important control variables such as the quality of schooling in relation to the quantity of schooling.

We use the framework of the augmented Solow model (Mankiw et al. 1992) to derive a specification that identifies an education externality within a production function framework. Our discussion aims to clarify how a macroeconomic return to education can be related to a Mincerian return (Mincer 1974), which has served in the literature as the standard benchmark for a microeconomic return to education. We show how a number of restrictions can be imposed on the specification in order to gain degrees of freedom for the estimation. Benchmark calculations based on a parameterization of our derived specification then suggest a range of plausible estimates for the education externality.

Based on these considerations, we argue that it will become difficult to estimate the exact size of the education externality when the difference between the macroeconomic and the microeconomic return to education is predicted to be in a narrow range of about 3 percentage points. Our empirical

results roughly confirm that the search for the size of the education externality comes close to the proverbial search for the needle in the haystack.

## 2. The Augmented Solow Model, Mincerian Schooling, and Education Externalities

### 2.1 *From the Mincer Equation to a Further Augmented Solow Model*

For more than 30 years, the Mincer equation (Mincer 1974) has been the workhorse for estimating the effect of schooling on individual wage income. The Mincer equation postulates a linear relation between the (log) wage as the dependent variable and education in the form of years of formal schooling and working experience, as in a specification like

$$\ln(w_i) = \beta_0 + \beta_1 school_i + \beta_2 exper_i, \quad (1)$$

where  $\ln(w_i)$  is the log wage of individual  $i$ ,  $school_i$  is individual years of schooling, and  $exper_i$  is individual years of working experience.<sup>1</sup> The coefficient  $\beta_1$  is the Mincerian private return to schooling. This coefficient indicates that an additional year of schooling raises the wage by  $\beta_1$  percent on average, all other things constant. Estimating the Mincer specification (1) with micro data, numerous studies for rich and poor countries suggest an average private return to schooling in the range of 6-10 percent, with the lower end as a representative estimate for an OECD average and the higher end as a representative estimate for a world average.<sup>2</sup>

Microeconomic estimates mainly focus on measuring the effect of an additional year of schooling on individual wage income and thus do not account for potential macroeconomic education externalities. However, the microeconomic relation between individual years of schooling and wages can be translated into a relation between average years of schooling and *per capita income*. The reason

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<sup>1</sup> The standard Mincer specification also includes a quadratic experience term, and may include further control variables.

<sup>2</sup> For instance, see Card (1999) or Hartog et al. (1999). For a recent summary of the empirical literature, see Psacharopoulos and Patrinos (2002).

is that the labor share (the share of total wages in GDP) does not follow a trend, neither over time (Piketty and Saez 2003) nor across countries (Bernanke and Gurkaynak 2001; Gollien 2002). But if labor shares are constant, wages are proportional to per-capita income. Hence the private return to schooling has to be multiplied by the labor share to derive a benchmark for the macroeconomic return to schooling that would be expected in the absence of an education externality. Put differently, a positive difference between an observed macro return to schooling and the *proportionally adjusted* private return to schooling would indicate the presence (and size) of an education externality.

The production function framework of the augmented Solow Model (Mankiw et al. 1992) provides a convenient possibility to clarify these considerations. For instance, one may start with a Cobb Douglas production function with physical capital  $K$  and labor  $L$  as the two basic inputs and assume that labor input is conditioned for the average level of education, such that

$$Y_j = B_j K_j^\alpha (h_j L_j)^{(1-\alpha)}, \quad (2)$$

where  $Y_j$  is the aggregate output of country  $j$ ,  $B_j$  is an index of the level of technology that is exogenous to individual firms within countries,  $K_j$  is the stock of aggregate physical capital,  $h_j$  is the average level of education, and  $L_j$  is the labor force. Educational capital accumulation may generate an externality, such that

$$B_j = A_j h_j^\gamma, \quad (3)$$

where  $A_j$  is an index of country-specific technology that grows over time  $t$  at an exogenous rate  $g$  with  $A_j = A(0)_j e^{g \cdot t}$  and  $\gamma$  represents the education externality.<sup>3</sup> The Mincer equation (1) suggests that the *average* level of education may be specified as a function of average schooling years and average experience (Bils and Klenow 2000), such that

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<sup>3</sup> The notation of the externality follows Lucas (1990).

$$h_j = \mu e^{\beta_1 school_j + \beta_2 exper_j}, \quad (4)$$

where  $\mu$  is a constant,  $school_j$  is average years of schooling in country  $j$  and  $exper_j$  is average years of working experience, and  $\beta_1$  and  $\beta_2$  represent the *average* individual private returns to schooling and experience. Substituting (4) and (3) into (2), it follows that

$$Y_j = A_j K_j^\alpha \left( \mu e^{\beta_1 school_j + \beta_2 exper_j} \right)^{(1-\alpha+\gamma)} (L_j)^{(1-\alpha)}. \quad (5)$$

Taking logs, dividing by  $L_j$ , and rearranging terms gives

$$\ln(y_j) = \ln A_j + (1-\alpha+\gamma) \ln \mu + \alpha \ln k_j + (1-\alpha+\gamma)(\beta_1 school_j + \beta_2 exper_j), \quad (6)$$

where  $y_j$  is GDP per person in country  $j$ , and  $k_j$  is capital per person in the labor force. Assuming that a constant fraction of output is saved and invested, and defining  $\tilde{k}_j = K/AL$  as the stock of capital per effective unit of labor and  $\tilde{y}_j = Y/AL$  as the level of output per effective unit of labor, it follows that the evolution of  $\tilde{k}_j$  over time is governed by

$$\begin{aligned} d\tilde{k}_j/dt &= s_j \cdot \tilde{y}_j - (n_j + g + \delta) \tilde{k}_j \\ &= s_j \cdot \tilde{k}_j^\alpha - (n_j + g + \delta) \tilde{k}_j \end{aligned} \quad (7)$$

where  $s_j = (I/GDP)_j$  is the share of investment in GDP,  $n_j$  is the growth rate of the labor force, and  $\delta$  is the rate of depreciation of physical capital. Mankiw et al. (1992) show that equation (7) implies that  $\tilde{k}_j$  converges to a steady state value  $\kappa_j$ , where  $d\tilde{k}_j/dt = 0$ :

$$\kappa_j = \left[ s_j / (n_j + g + \delta) \right]^{1/(1-\alpha)}. \quad (8)$$

Substituting the steady state capital intensity  $\kappa_j$  for  $k_j$  and taking logs, equation (6) can be rewritten as

$$\ln(y_j) = \text{const} + A_j + \frac{\alpha}{1-\alpha} \ln \left[ \frac{s_j}{n_j + g + \delta} \right] + (1-\alpha+\gamma)(\beta_1 \text{school}_j + \beta_2 \text{exper}_j), \quad (9)$$

where  $\text{const} = (1-\alpha+\gamma) \ln \mu$  is a constant.<sup>4</sup> Equation (9) predicts that the coefficient on the investment share equals in absolute value the coefficient on labor force growth (conditioned by  $g$  and  $\delta$ ). For a capital share in factor income of one third, the size of this coefficient is predicted to be  $(1/3)/(1-(1/3)) = 0.5$ . Furthermore, the coefficients of the two components of the average level of education, schooling years and experience, are predicted to be the same except for the difference in the two private rate of returns,  $\beta_1$  and  $\beta_2$ . Since the private returns to schooling and experience are known from a large microeconomic literature, the education externality  $\gamma$  remains as the only free parameter of equation (9).

Imposing on equation (9) the empirical restrictions on the capital share and on the two private returns, we can rearrange terms as

$$\ln(y_j) - \frac{\bar{\alpha}}{1-\bar{\alpha}} \ln \left[ \frac{s_j}{n_j + g + \delta} \right] = \text{const} + A_j + (1-\bar{\alpha}+\gamma)(\bar{\beta}_1 \text{school}_j + \bar{\beta}_2 \text{exper}_j), \quad (10)$$

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<sup>4</sup> For a general discussion of steady state values of one-sector growth models, see Jensen et al. (2005) and Richter (2007).



where the restricted parameters are indicated by a bar. In this specification, the coefficient to be estimated is  $(1 - \bar{\alpha} + \gamma)$ . In the absence of an education externality, the coefficient should equal two thirds for a capital share of one third.

## 2.2. *Macro Returns and the Size of the Education Externality: Benchmark Calculations*

Equation (10) can be used to derive benchmark values for the macroeconomic returns to schooling ( $\lambda_{school}$ ) and experience ( $\lambda_{exper}$ ):

$$\lambda_{school} = (1 - \bar{\alpha} + \gamma)\bar{\beta}_1 \quad \text{and} \quad (11a)$$

$$\lambda_{exper} = (1 - \bar{\alpha} + \gamma)\bar{\beta}_2 . \quad (11b)$$

Hence with a capital share in factor income of  $\bar{\alpha} = 1/3$  and with  $\bar{\beta}_1 = 0.1$  as an upper bound for the private return to schooling, we would expect to find a value of the macro-Mincer return to schooling of  $\lambda_{school} = 0.067$  if there is no education externality. Similarly, we would expect to find a value of the macro return to experience of  $\lambda_{exper} = 0.02$  for an average private return of 3 percent ( $\bar{\beta}_2 = 0.03$ ) if there is no education externality. That is, a macro return to schooling of 10 percent or a macro return to experience of 3 percent would indicate the presence of a large education externality.

Prevailing subsidies to public education are a major reason to expect empirical estimates of  $\lambda_{school}$  and  $\lambda_{exper}$  above the benchmark values of 6.7 percent and 2 percent. Given that the instructional costs of schooling are mainly covered by public expenditures in most countries, the question of interest is by *how much* the empirical estimates of  $\lambda_{school}$  and  $\lambda_{exper}$  can be expected to differ from the benchmark values. For instance in the United States, about 90 percent of the instructional costs at public colleges are paid by federal and state governments. Based on US data, Heckman and Klenow (1997) calculate that due to these subsidies, an individual will be indifferent between additional schooling (college) and working if the private Mincerian return to schooling is

about 9 percent, which is in line with the typical micro-Mincer estimates. From the point of view of the government, they calculate that the observed level of schooling subsidies is justified if the macroeconomic return to schooling is about 3 percentage points higher than the benchmark value for the private return. So based on our no-externality macro benchmark value of 6.7 percent, we would expect to find a macro return to schooling of about 10 percent in order to justify present levels of schooling subsidies.

A similar benchmark follows from a calculation that considers the macroeconomic return to physical capital. If capital earns its marginal product ( $MPK$ ), it can be measured as  $MPK = \bar{\alpha} / (K/Y)$ . With a capital share of one third and a capital output ratio of 3, as discussed by Mankiw et al. (1992) for the US, it follows that the implied macroeconomic return to physical capital equals 11 percent. With a lower capital output ratio of 2, which represents an average value for rich and poor countries, the implied macroeconomic return to capital would rise to about 17 percent. For a lower bound of the capital share of  $\bar{\alpha} = 1/4$ , the macroeconomic return to capital would range from 8 percent to 13 percent, conditional on the applied capital output ratio. Since the macroeconomic returns to physical and educational capital investments will tend to be equalized, we consider a macro return to schooling of about 10 percent as a lower bound.

Given that the level of schooling subsidies reaches similar levels as in the US in most countries around the world, an estimated macro return to schooling in the range of 10 percent would imply that a large education externality is already internalized by the public subsidization of schooling. For instance, equation (11a) suggests a benchmark value for the education externality  $\gamma$  of one third for a macro return of 10 percent, a micro return of 10 percent, and a capital share of one third.<sup>5</sup>

Equation (11a) also suggests that an estimate of an education externality that differs in a statistically significant way from the benchmark estimate will be hard to come by. This is because even small and statistically insignificant differences in macro returns to schooling imply large differences in the size of the education externality. For all other parameters as before, an estimate of

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<sup>5</sup> Lucas (1990) derived a comparable education externality of 0.4 on the basis of US data for a capital share of 1/4 and the *growth rates* of per capita income and average years of schooling.

the macro return of 15 percent would already imply an education externality of 0.8, which looks implausibly high within a production function framework.<sup>6</sup> Thus a priori it seems that empirical analyses can at best reveal whether there is a positive education externality at all, but not by how much such an externality may differ from a benchmark value of one third that considers the prevailing level of education subsidies.

### 3. Empirical Results

#### 3.1 Main Variables and Measures of Schooling Quality

We begin our empirical analysis with a basic sample of 82<sup>7</sup> countries for which we have data on the variables specified in equation (9), mainly for the base year 1995. Our data on GDP per worker ( $y$ ), investment ( $s = I / GDP$ ), and labor force growth ( $n$ ) are taken from PWT 6.1 (Heston et al. 2002). We assume a constant rate  $g = 0.02$  and a constant rate  $\delta = 0.05$ , as is standard in the related literature (see, e.g., Barro et al. 1995). Average years of schooling (*school*) are taken from Barro and Lee (2001) and our measure of the average years of experience of the workforce (*exper*) is based on data taken from UN (1997). The appendix A.1 includes detailed information on the definitions and sources of all variables used in the empirical analysis.

Our preferred measure of schooling years takes into account that the *quality* of schooling is not the same across countries. Recent comparative studies of international educational standards such as the PISA study (OECD 2001) or previously the TIMS study (TIMSS 1996) have shown that there are large international differences in the quality of schooling. We address this problem by correcting the observed average years of schooling with a measure of schooling quality.

Hanushek and Kimko (2000) report standardized test scores in international comparisons of student achievement in "percent correct" format. For countries not covered by Hanushek and Kimko (2000) but included in our sample, we have imputed the missing values as averages of region- and

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<sup>6</sup> An education externality of 0.8 would imply that a doubling of all inputs would increase output by a factor of 3.5.

<sup>7</sup> See appendix A.2.

income-based estimates, as reported in Gundlach et al. (2002). We then transform the "percent correct" format into a format with an average of 500 and a standard deviation of 100 by solving for  $a$  and  $b$  in:

$$average(percent\ correct) * b + a = 500, \quad \text{and} \quad (12)$$

$$std(percent\ correct) * b = 100. \quad (13)$$

The format with an average of 500 and a standard deviation of 100 has been established as a standard reporting format in recent international studies of student achievement (OECD 2001, TIMSS 1996). According to these studies, a test score difference of about 50 points represents an average performance difference of about one year of schooling (one grade). Regressing average years of schooling on our transformed test score measure, we find that one year of schooling in the population aged 15-65 is associated with a test score difference of about 60 points, close to the results reported for TIMSS and PISA. We define our measure of quality-adjusted schooling as

$$qschool_j = school_j + (test_j - 500)/60, \quad (14)$$

where  $test_j$  is the transformed test score result for country  $j$ . That is, for countries with test score results close to the international average of 500 points (like the United States), the quality adjustment does not make an important difference, but a country with a test score result of about 380 points (like Kenya) would have a quality-adjusted measure of schooling that is about two years below the average years of schooling reported by Barro and Lee (2001).

### 3.2 Empirical Estimates of the Education Externality

Our empirical analysis mainly aims at estimating the sign of the presumed education externality, and more ambitiously at estimating its size. As discussed in Section 2, we impose as restrictions a capital share of  $\bar{\alpha} = 1/3$ , and private returns to schooling of  $\bar{\beta}_1 = 0.1$  and to experience of  $\bar{\beta}_2 = 0.03$ . By imposing the empirical restrictions we can circumvent the problem that otherwise the parameters of

equation (9) cannot be estimated simultaneously because we have the four potentially unknown parameters  $\alpha$ ,  $\gamma$ ,  $\beta_1$ , and  $\beta_2$ , but only the three explanatory variables (conditional) investment ( $\ln s_j - \ln(n_j + g + \delta)$ ), average years of schooling with or without quality adjustment ( $qschool$ ,  $school$ ), and experience ( $exper$ ). So we estimate the education externality  $\gamma$  with the restricted equation (10). However, equation (10) suggests that we should control for international differences in the level of technology  $A_j$ .

In appendix A.3 we document that a number of variables that have been suggested in the literature as proxies for differences in technology across countries are closely correlated with a set of continental dummies for Sub-Saharan Africa (*SSA*), Asia (*ASIA*), and Latin America (*LATINAM*). For instance, we find that this set of three continental dummies can explain between 40 and 70 percent of the variation of each of the following variables, which all may reflect certain aspects of international differences in a broad concept of technology: a measure of life expectancy (Heckman and Klenow 1997), two alternative measures of the prevalence of malaria (Sachs 2003), and four alternative measures of institutional quality (Hall and Jones 1999, Acemoglu et al. 2001, Rodrik et al. 2004). These empirical results motivate our approach to control for international differences in a broad concept of technology with three continental dummies.

We get three basic results. We obtain a statistically significant education externality with a large point estimate of 2.6 if we do not control for international differences in the quality of schooling and in investment and technology (Table 1, column (1)). From an economic perspective, i.e., relative to our benchmark value of 0.33 derived in section 2, an estimated externality of 2.6 appears to reflect an omitted variables bias, since measures of education and experience are likely to be positively correlated with the quality of schooling, the level of investment, and the level of technology. Hence the estimated externality may be biased upwardly.

Table 1 – Estimating the Education Externality for Alternative Restrictions

	Dependent variable: Log real GDP per worker in 1995			
	(1)	(2)	(3)	(4)
$\bar{\beta}_1 school_j + \bar{\beta}_2 exper_j$	3.26 (0.197)		1.51 (0.249)	
$\bar{\beta}_1 qschool_j + \bar{\beta}_2 exper_j$		1.97 (0.154)		0.77 (0.152)
$\ln(s_j / (n_j + g + \delta))$			0.5 (imposed)	0.5 (imposed)
<i>SSA</i>			-0.97 (0.194)	-1.23 (0.181)
<i>ASIA</i>			-0.49 (0.168)	-0.65 (0.168)
<i>LATINAM</i>			-0.20 (0.143)	-0.29 (0.147)
$\gamma$	2.60 (0.197) [0.000]	1.30 (0.154) [0.000]	0.85 (0.249) [0.001]	0.11 (0.152) [0.239]
$\bar{R}^2$	0.771	0.668	0.758	0.733
No. of observations (countries)	82	82	82	82

Notes: We assume  $\bar{\alpha} = 1/3$ ,  $\bar{\beta}_1 = 0.1$ , and  $\bar{\beta}_2 = 0.03$ . The capital externality parameter  $\gamma$  is calculated as the reported regression coefficient on the combined measure of education minus  $(1 - \bar{\alpha}) = 2/3$ , see equation (10). Standard errors in parentheses;  $p$ -values of the one-sided hypothesis  $\gamma = 0$  against  $\gamma > 0$  in brackets.

Once we control for differences in the quality of schooling (column (2)) *or* for differences in the levels of investment and technology (column (3)), the point estimates of the externality are substantially reduced and remain statistically significant, but they are still too high to be plausible within the production function framework outlined in section 2.

Our preferred specification controls for differences in the quality of schooling *and* for differences in the levels of investment and technology (column (4)). Here we find that the point estimate of the education externality drops to 0.11 and becomes statistically insignificant. Even the one-sided hypothesis of  $\gamma = 0$  against  $\gamma > 0$  cannot be rejected at a  $p$ -value of 0.20. So on the basis of our OLS estimates, we would conclude that the education externality is found to be too large when

the empirical model is not properly restricted, and found to be absent when all presumed control variables are included. Taken at face value, this result would suggest that the observed level of subsidization of schooling is not justified for reasons of economic efficiency.

One caveat of our results is that they may be biased because of endogenous regressors and especially because of measurement error. For instance, schooling and experience may be imperfect proxies for a true measure of educational capital, with or without an adjustment for schooling quality. To control for this potential sources of downward bias, one would have to consider instrumental variables estimation (IVs). We leave a detailed discussion of IV estimation for further research.

#### **4. Summary**

Our paper shows how the augmented Solow model can be combined with the Mincer equation to derive a simple production function specification that allows us to estimate an education externality. With this framework, we can show how the well-known private rates of return to schooling and experience relate to a macroeconomic rate of return that takes into account all external costs and benefits of educational capital accumulation. For reasonable parameter values, we can also show that a macroeconomic return that exceeds an (adjusted) private benchmark return by a few percentage points implies a large difference for the education externality.

These considerations reveal why previous empirical studies have not succeeded in providing a clear-cut picture. With cross-country data, even a five percentage point difference in macroeconomic and private rates of return to schooling may not be large enough to allow for a statistically significant estimate of the size of the education externality, and may often not even suffice to identify its sign. Our OLS regression results confirm that what may appear as a statistically significant estimate of an education externality disappears once the underlying specification is properly restricted. It remains to be seen whether IV estimation results may provide a different answer.

## Appendix

### A 1. Definitions and Sources of Variables

<i>A</i>	Level of technology, approximated by continental dummies.
<i>ASIA</i>	Dummy variable for Asian countries; excludes Japan, includes Papua NG and other Oceania.
$\delta$	Depreciation rate (of physical capital); assumed to be 5 percent.
<i>exper</i>	Average years of experience of the working age population in 1990, calculated as indicated in Heckman and Klenow (1997). Based on the age distribution of the population aged 5-64 years and on school enrollment rates, which can be combined to calculate the share of the population aged 5-64 years that is working (not in school). For each age group, work experience is calculated as age <i>minus</i> schooling attainment <i>minus</i> 6 years. Our measure is the working-population-weighted average of these experience levels. Source: UN (1997); Barro and Lee (2001); own calculations.
<i>exprop</i>	Index of protection against expropriation, constructed by Political Risk Services. Source: taken from McArthur and Sachs (2001).
<i>g</i>	Rate of technological change; assumed to be 2 percent.
<i>gadp</i>	Index of government anti-diversion policies; calculated as an unweighted average of five variables: law and order, bureaucratic quality, corruption, risk of expropriation, and government repudiation of contracts. Source: Hall and Jones (1999).
<i>kaufm</i>	Unweighted average governance indicator (1996), based on six aggregated survey measures: voice and accountability, political stability, government effectiveness, regulatory quality, rule of law, and control of corruption. Source: Kaufmann et al. (2004).
<i>LATINAM</i>	Dummy variable for Latin American countries.
<i>life</i>	Life expectancy at birth, 1990. Source: World Bank (2005).
<i>malrisk</i>	Proportion of each country's population at risk of malaria transmission, involving three largely non-fatal species of the malaria pathogen ( <i>Plasmodium vivax</i> , <i>P. malariae</i> , <i>P. ovale</i> ) (1994); measured on a [0,1] scale; dataset as of 27 October 2003. Source: Sachs (2003), here taken from <a href="http://www.earth.columbia.edu/about/director/-malaria/index.html#datasets">http://www.earth.columbia.edu/about/director/-malaria/index.html#datasets</a> .
<i>mfalrisk</i>	Proportion of a country's population at risk of falciparum malaria transmission (1994); measured on a [0,1] scale. Revised version, dataset as of 27 October 2003. Source: Sachs (2003), here taken from <a href="http://www.earth.columbia.edu/about/-director/malaria/index.html#datasets">http://www.earth.columbia.edu/about/-director/malaria/index.html#datasets</a> .
<i>n</i>	Average growth rate of the workforce in 1960-1995. Source: Heston et al. (2002); own calculations.



<i>school</i>	Average years of schooling of the population aged 15 to 65 in 1995. Source: Barro and Lee (2001)
<i>qschool</i>	Quality-adjusted average years of schooling. For an explanation see text. Source: Hanushek and Kimko (2000), Barro and Lee (2001), Gundlach et al. (2002); own calculations.
<i>s</i>	Average share of investment in real GDP ( $I/GDP$ ) in 1960-1995. Source: Heston et al. (2002); own calculations.
<i>socinf</i>	Index of social infrastructure based on the years that a country is open to international trade ( <i>yrsopen</i> ) and the measure <i>gdp</i> ; calculated as $(gdp+yrsopen)/2$ . Source: Hall and Jones (1999).
<i>SSA</i>	Dummy variable for Sub-Saharan African countries.
<i>y</i>	Real GDP per worker in 1995. Source: Heston et al. (2002).

## A 2. Sample of Countries

Our sample is determined by data availability and includes 82 countries (see Table 1). These countries are:

Argentina; Australia; Austria, Bangladesh, Barbados, Belgium, Benin, Bolivia, Botswana, Brazil, Cameroon, Canada, Central African Republic, Chile, China, Colombia, Congo, Costa Rica, Cyprus, Denmark, Dominican Republic, Ecuador, Egypt, El Salvador, Fiji, Finland, France, Gambia, Germany West, Ghana, Greece, Guyana, Haiti, Honduras, Hong Kong, Hungary, Iceland, India, Indonesia, Iran, Ireland, Israel, Italy, Jamaica, Japan, Kenya, Mali, Mexico, Netherlands, New Zealand, Nicaragua, Niger, Panama, Papua New Guinea, Paraguay, Peru, Philippines, Poland, Portugal, Rwanda, Senegal, Sierra Leone, Singapore, South Africa, Spain, Sri Lanka, Sweden, Switzerland, Syrian Arab Republic, Tanzania, Thailand, Togo, Trinidad and Tobago, Tunisia, Turkey, Uganda, Uruguay, United Kingdom, United States, Venezuela, Zambia, Zimbabwe.

### A.3 Appendix Table – Continental Dummies as Determinants of Technology

Proxy of technology (Dependent variable)	Explanatory variables			$\bar{R}^2$	No. of obs.
	<i>SSA</i>	<i>ASIA</i>	<i>LATINAM</i>		
Life expectancy <i>life</i>	-24.04 (1.66)	-8.44 (1.99)	-5.93 (1.59)	0.72	82
Disease ecology <i>mfalrisk</i>	0.87 (0.06)	0.28 (0.07)	0.10 (0.06)	0.75	80
<i>malrisk</i>	0.88 (0.08)	0.44 (0.10)	0.24 (0.07)	0.62	80
Institutional quality <i>kaufm</i>	-1.43 (0.19)	-0.88 (0.22)	-1.03 (0.17)	0.48	77
<i>socinf</i>	-0.44 (0.06)	-0.22 (0.07)	-0.31 (0.05)	0.43	82
<i>gadp</i>	-0.34 (0.04)	-0.25 (0.05)	0.31 (0.04)	0.52	82
<i>exprop</i>	-2.74 (0.40)	-1.61 (0.47)	-2.41 (0.37)	0.46	74

Note: All regressions include a constant term; standard errors in parenthesis (OLS). For details see text. Definitions of variables and sources of data are given in appendix A.1.

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