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Zentrum für Europäische Integrationsforschung
Center for European Integration Studies
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Euribor Futures Market:
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ECB Policy Announcements**

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The Euribor Futures Market: Efficiency and the Impact of ECB Policy Announcements*

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Abstract

For an effective and smooth monetary policy, it is important that interest rate expectations are in line with central bank policy intentions. The predictability of money market interest rates is, therefore, an indicator of transparency and clarity in the communication of monetary policy and of the effectiveness of monetary policy implementation.

In this paper, we analyse three aspects of the predictability of money market rates in the European Monetary Union (EMU). The first is the efficiency of the three-month Euribor interest rate futures markets. The second aspect is the effect of ECB policy announcements on the volatility of Euribor futures rates, and the third aspect is the effect of ECB policy announcements on the prediction error contained in Euribor futures rates. We find that the new Euro money markets were able to predict short-term rates well. Our results suggest that the ECB communication of monetary policy has worked well during the first years of EMU and that the predictability of ECB policy decisions seems to have improved over

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time. ECB Council decisions still cause some surprises, but their effect on volatility is small.

I. Introduction

Short-term money market rates play a key role in the European Central Bank's (ECB) conduct of monetary policy. Since current money market rates are determined by the markets' expectations about future policy rates, it is important for an effective and smooth monetary policy that interest rate expectations are in line with central bank policy intentions (Poole and Rasche 2000). The predictability of money market interest rates is, therefore, an indicator of transparency and clarity in the communication of monetary policy and of the effectiveness of monetary policy implementation.

The predictability of money market rates has recently been investigated by a number of empirical studies. One approach starts from the rational expectations' hypothesis of the term structure of interest rates, which holds that money market rates of longer maturities should be unbiased and efficient predictors of future short-term rates. Using US data, Mankiw and Miron (1986), Rudebusch (1995), Balduzzi et al. (1997), Favero and Mosca (2001) test this hypothesis for various interest rates and analyse whether the term structure serves as a predictor of Federal Reserve funds target changes. The general finding is that the process of Federal Reserve policy moves is not well anticipated by the market, but Favero and Mosca (2001) conclude that uncertainty about monetary policy has been reduced from 1994 onwards.

Another approach is based on an application of the efficient market hypothesis (EMH) (Fama 1970) to the market for interest rate futures. Efficiency of this market would indicate that market participants are able, on average, to predict future spot rates correctly. This would point to a well-functioning communications policy of the central bank regarding its intended interest rate policy. Cole et al. (1991), Cole and Reichenstein (1994) and Krueger and Kuttner (1996) test for the efficiency of different short-term interest rate futures markets. Their empirical results generally support the EMH.

A third approach focuses on the effects of monetary policy announcements. If markets are efficient, a systematic response of asset prices or interest rates to central bank policy announcements would indicate that these announcements cause market participants to revise their expectations; that is, that they contain some 'news'. Depending on the size of revisions, this could imply that the central bank did not fully succeed in a smooth management of information regarding its monetary policy. Cook and Hahn (1989), Roley and Sellon (1998), Thornton (1998), Bomfim and Reinhart

(2000), Poole and Rasche (2000), Bomfim (2003), Rigobon and Sack (2001) and Kuttner (2001) find that monetary policy announcements by the Federal Reserve do have significant effects on short-term interest rates and stock prices.

This general approach has recently also been applied to the new Euro money market. The first paper to do this, Hartmann et al. (2001), uses intraday overnight interbank rates from voice brokers and electronic trading in five euro-area countries to estimate the prediction error of money market operators around ECB monetary policy decisions. Another paper, by Gaspar et al. (2002), examines the effect of ECB policy on the level and volatility of the daily overnight interbank rate (EONIA) using a GARCH model. Both papers conclude that market participants were able to predict ECB monetary policy decisions relatively well. In contrast, the pamphlet by Ross (2002) – applying different approaches to the euro area, the USA and the UK – argues that the ECB is less predictable than the Federal Reserve or the Bank of England.

In this paper, we add a new perspective to this debate. We analyse three aspects of the predictability of money market rates in the new market environment of the European Monetary Union (EMU). The first is the efficiency of the Euribor interest rate futures markets. We focus on the three-month Euribor future, which is the most actively traded interest rate derivative in the European money market. The second aspect is the effect of ECB policy announcements on the volatility of Euribor futures rates. The third aspect is the effect of ECB policy announcements on the prediction error contained in Euribor futures rates. A systematic decrease in the absolute prediction error would indicate that the announcement released by the ECB on a Governing Council day contains new information and, therefore, improves the markets' ability to predict future spot rates. This, too, would imply less successful information management by the ECB.

The Euribor market is of special interest, because it is a new market that emerged only with the start of EMU in 1999.¹ Our paper is the first to analyse the performance of futures rates in this market. Given the small number of futures contracts traded since its start, we use an efficiency test based on a panel-data approach instead of the conventional time-series approach. By using data since trading in Euribor futures first began in December 1998, we can check whether the markets' ability to predict ECB interest rate decisions has improved over the first years of EMU, and how the ECB's information management has developed over time.

¹For characterizations of the euro money market see Hartmann et al. (2001), and Ewerhart et al. (2003).

The main results of this paper can be summarized as follows. First, we find that Euribor futures rates are indeed unbiased predictors of future spot rates. Furthermore, futures rates with forecast horizons up to four months are also (weakly) informationally efficient. This suggests that ECB policy decisions have, on average, been predictable and that, considering the fact that it is a new central bank operating in a new monetary environment, the ECB's information policy has worked surprisingly well.

Second, we find that the volatility of futures rates is significantly higher on days when the ECB Governing Council met than on non-Council days. At a closer look, however, it seems that market participants anticipated the majority of the ECB's policy decisions correctly. Only a few Council decisions seem to have taken the market by surprise. Third, we find that ECB policy announcements have no systematic effect on the absolute size of the prediction error contained in Euribor futures rates. These results corroborate our conclusion that, with few exceptions, the ECB's information policy has contributed to a smooth operation of monetary policy.

The paper proceeds as follows. Section II develops our test of market efficiency and presents the results. Section III presents our analysis of the effects of ECB policy announcements on Euribor futures rates. Section IV concludes.

II. Efficiency of the Euribor Futures Market

A. Testing Money Market Efficiency: A Panel Approach

Let r_t be the spot interest rate on a three-month interbank deposit at time t and $f_{i,t}$ the futures rate at time $t-i$ for a futures contract expiring at time t that has the same deposit as its underlying asset. Assuming that investors are risk neutral, arbitrage requires that the futures rate $f_{i,t}$ be equal to the spot rate r_t expected at time $t-i$, $E_{t-i}(r_t)$, where $E_{t-i}(\cdot)$ is the conditional expectation given all information available to market participants at time $t-i$. The EMH holds that market participants incorporate all relevant and available information in the determination of futures rates, and that, therefore, any difference between the spot rate, r_t , and the futures rate, $f_{i,t}$, is unpredictable given all information available at time $t-i$. Thus, $r_t = f_{i,t} + \varepsilon_{t,i}$ with $E_{t-i}(\varepsilon_{t,i}) = 0$.

The standard way of testing for the unbiasedness and efficiency of interest rate futures markets in the literature is to run a time-series regression of the following type:²

$$r_t = \alpha_i + \beta_i f_{i,t} + \gamma_i X_{i,t} + \varepsilon_{i,t} \quad (1)$$

²Compare, for example, Cole et al. (1991), Krehbiel and Adkins (1994), Cole and Reichenstein (1994), Krueger and Kuttner (1996) and Cuthbertson (1996).

with $t = 1, \dots, T$. α_i can be interpreted as a risk premium, X_{it} is a column vector of variables contained in the set of information available to the market at time $t-i$, and γ_i is a corresponding row vector. Based on equation (1), unbiasedness and efficiency requires that $\alpha_i = \beta_i - 1 = \gamma_i = 0$ and that the expectation error, $\varepsilon_{i,t}$ is serially uncorrelated.

In this standard test, the forecast horizon, i , is fixed and the data are non-overlapping, that is the distance between any two observations r_t and r_{t-1} is at least of length i . The test thus uses one futures rate quoted before the settlement day for each futures contract. This requires a sufficiently large number of contracts observed in the market. For the Euribor market, this approach is not practicable, simply because not enough contracts have been traded since the market started in December 1998. In view of this, we propose a test based on a panel approach similar to Dunis and Keller (1995), who perform an efficiency test for different currency option markets. The panel approach exploits the fact that futures rates are quoted on a daily basis, so that, for a contract expiring at time t , multiple futures rates with different times to expiration can be observed.

Let T be the number of contracts included in the sample and N the number of futures rates observed for each contract. For each contract $t = 1, \dots, T$, we observe one spot rate, r_t , on the settlement day, and N futures rates, $f_{i,t}$, $i = 1, \dots, N$. Thus, we build a panel forming N groups of futures rates with the same forecast horizon i , where each group contains T observations. The panel estimator used below takes the forecast horizon i as the cross-section and the settlement date as the time-series dimension. A necessary condition for this approach to be feasible is the poolability of futures rates with different forecast horizons, i.e., $\beta_i = \beta$, and $\gamma_i = \gamma$ for all $i = 1, \dots, N$ (Baltagi 1995). If poolability holds, equation (1) can be rewritten as

$$r_t = \alpha_i + \beta f_{i,t} + \gamma X_{i,t} + \varepsilon_{i,t} \quad (2)$$

with $t = 1, \dots, T$; $i = 1, \dots, N$. The panel set-up thus gives us NT observations to test for market efficiency. Note that the panel allows for risk premia, α_i , to vary with the length of the forecast horizon, as suggested by Fama (1984). The null hypothesis of efficiency and unbiasedness is then expressed by the condition that $\alpha_i = \beta - 1 = \gamma = 0$ for all i , and that the expectation error $\varepsilon_{i,t}$ is serially uncorrelated.

Let $\Omega_{i,t}$ and $\Omega_{j,t}$ be the information sets underlying the forecasts of the spot market rate r_t embedded in two different futures rates, $f_{i,t}$ and $f_{j,t}$, with $0 < j < i \leq N$. Our panel structure implies that $\Omega_{i,t}$ is contained in $\Omega_{j,t}$, $\Omega_{i,t} \subset \Omega_{j,t}$ and, therefore, the error terms $\varepsilon_{i,t}$ and $\varepsilon_{j,t}$ are correlated. Furthermore, the variance of $\varepsilon_{j,t}$ can be expected to be smaller than the variance of $\varepsilon_{i,t}$. To account for this, we use an OLS estimator with panel-corrected standard errors proposed by Beck and Katz (1995), which corrects

for heteroskedasticity, contemporaneous correlation across groups and, if necessary, serial correlation.

B. Data

A Euribor future is a futures contract with a Euribor deposit as the underlying asset. Since 1 January 1999, the Euribor has been used as the European money market reference rate for the unsecured market. One-month and three-month Euribor futures have been traded on the derivatives market since December 1998. According to information from the LIFFE (London International Financial Futures and Options Exchange), the three-month Euribor futures contract currently accounts for over 99% of the euro denominated short-term interest rate derivatives market. During the third quarter of 2003, 34.6 million Euribor futures contracts were traded, an increase of 16% compared to the equivalent period the year before.

The three-month Euribor future is a commitment to engage in a three month loan or deposit of a face value of 1000,000 Euros. There are four delivery dates during a year, namely the third Wednesday of March, June, September and December. The last trading day of each futures contract is two trading days prior to the relevant settlement day. Futures prices are quoted on a daily basis and the interest rate contracted equals 100 less the futures price.

The data used in this study are provided by LIFFE.³ Our sample contains daily closing rates of the 19 three-month Euribor futures contracts that settled between March 1999 and September 2003, i.e. $t = 1, \dots, 19$. We use all futures rates with forecast horizons of up to six months for each futures contract, i.e. $i = 1, \dots, 183$. After accounting for weekends, this yields a total of 131 cross-sections with 19 observations each. The first futures rate in our data sample was priced on 15 December 1998 and the last on 15 September 2003. A Chow test does not reject the null hypothesis of poolability. This means that the panel approach is feasible given our data.⁴

Figures 1–6 plot the futures rates and the corresponding spot interest rates with different forecast horizons.⁵ Financial markets over-predict future spot rates, when the futures rate is above the corresponding spot rate, and they

³Data available at <http://www.liffe-style.com>.

⁴The results are available from the authors on request.

⁵Note that all futures rates of a futures contract predict the same spot rate realized at the third Wednesday of the relevant month (March, June, September, December); hence the figure shows the spot rate moving in steps and the futures rates fluctuating around it.

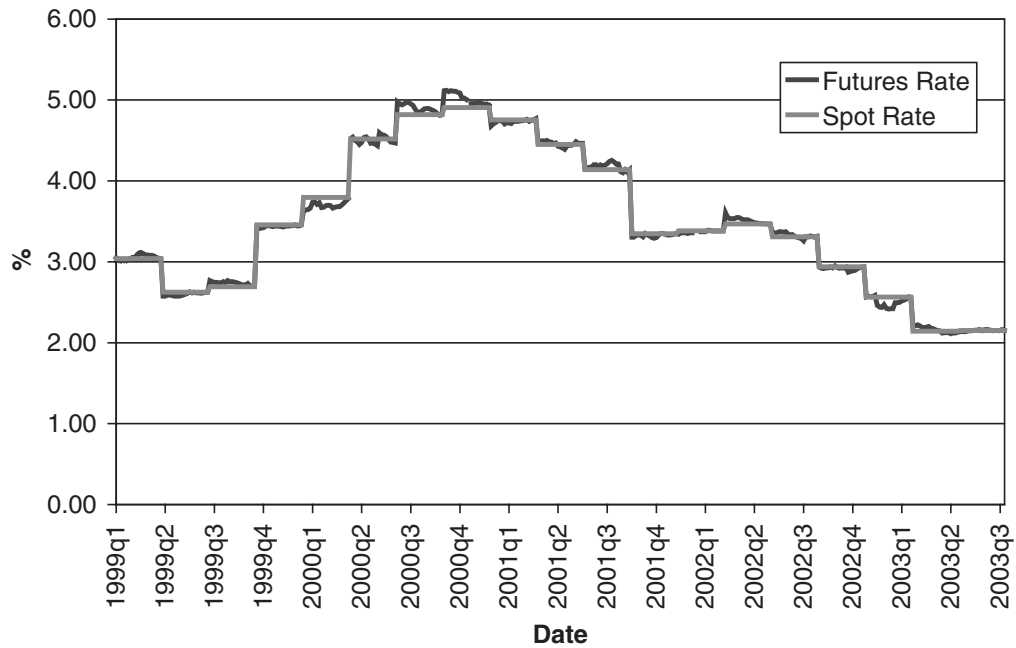


Figure 1: Futures Rates with Forecast Horizon of 1–31 days

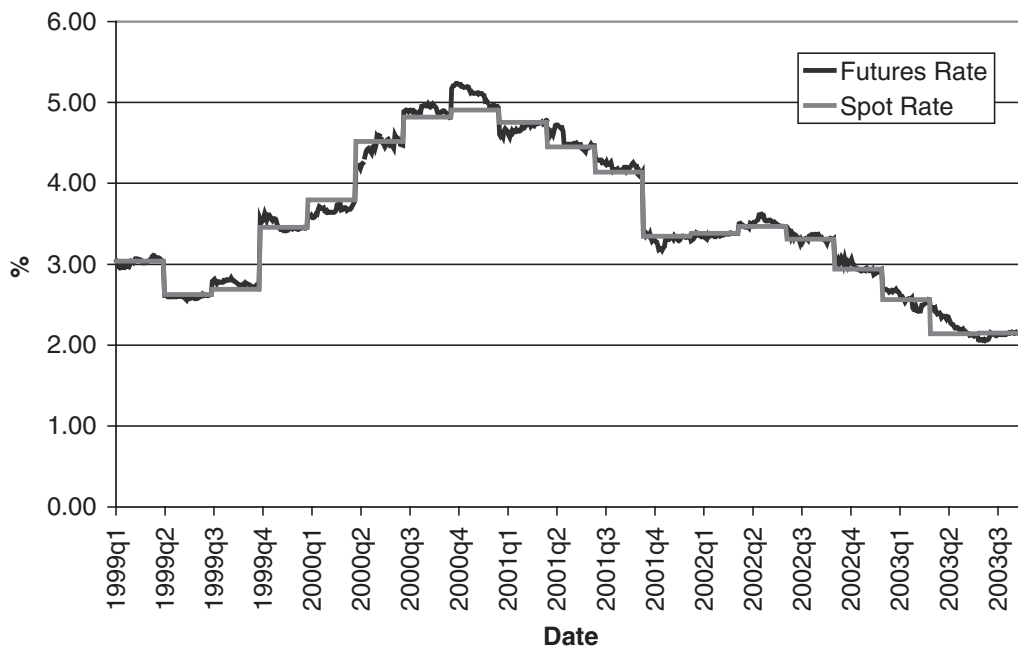


Figure 2: Futures Rates with Forecast Horizon of 1–61 days

under-predict future spot rates, when the futures rate is below that. The figures show that the Euribor futures rates are generally close to the corresponding spot rates, especially when their settlement dates approach. A comparison between Figure 1 and Figure 6 shows that, unsurprisingly, the ability of futures rates to predict spot interest rates decreases with the length

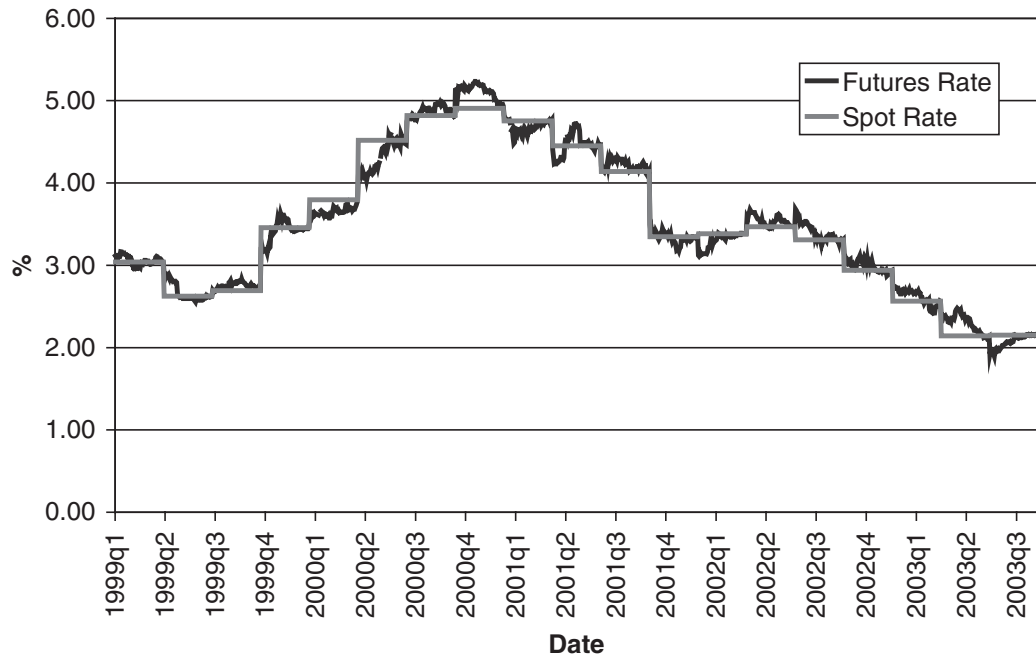


Figure 3: Futures Rates with Forecast Horizon of 1–91 days



Figure 4: Futures Rates with Forecast Horizon of 1–122 days

of the forecast horizon. The figures suggest that markets tend to under-predict interest rates in times, when spot rates follow an upward trend, and they tend to over-predict interest rates, when there is a downward trend of spot interest rates. For example, Figure 6 shows that the Euribor futures rates of the June 2002 and September 2002 contracts contained positive prediction

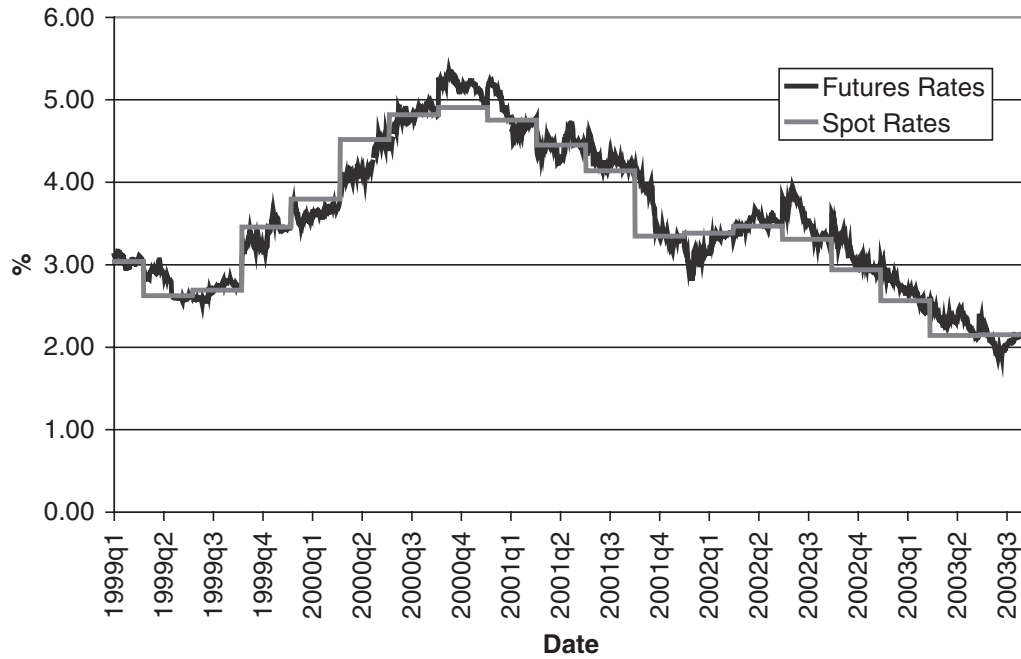


Figure 5: Futures Rates with Forecast Horizon of 1–153 days

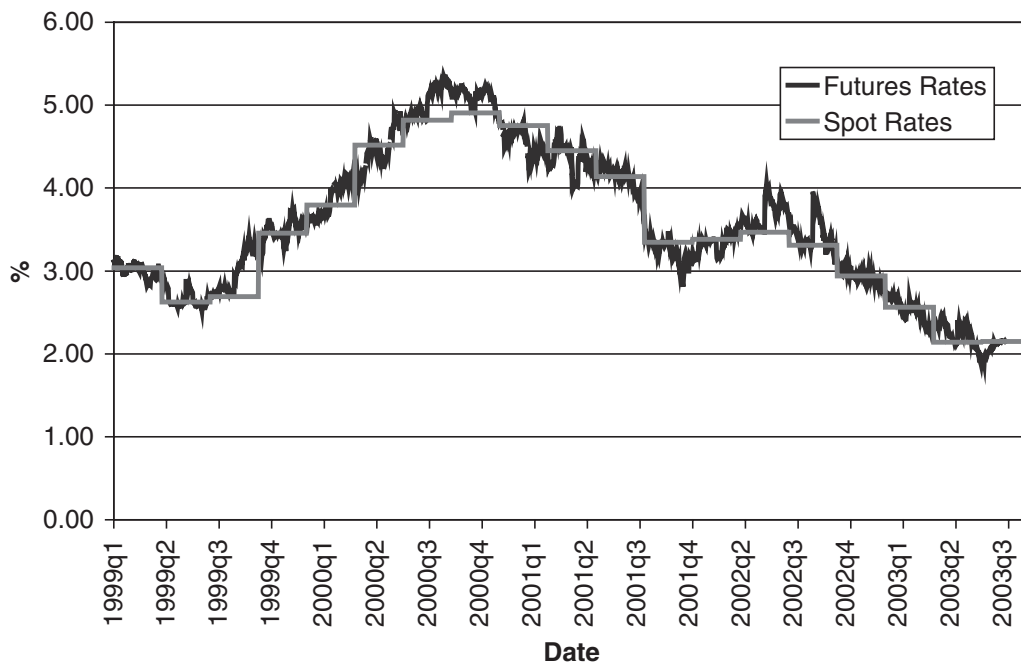


Figure 6: Futures Rates with Forecast Horizon of 1–183 days

errors of almost 50 basis points. Thus, market participants expected a rise of the Euribor interest rate, which actually did not happen. Contrarily, the futures rates of the March 2002 contract under-predicted future spot rates by around 50 basis points. Market participants apparently expected a

continuation of the downward trend of Euribor interest rates, but actual spot rates slightly increased.

C. Results

It is plausible that the markets' forecasting ability diminishes as the time to expiration of the futures rates i increases.⁶ Therefore, the results of the efficiency and unbiasedness tests may depend on the length of the forecast period. In view of this, we perform our tests for forecasting horizons of increasing length: one month ($i = 1, \dots, 31$), two months ($i = 1, \dots, 61$), three months ($i = 1, \dots, 91$), four months ($i = 1, \dots, 122$), five months ($i = 1, \dots, 153$) and finally six months ($i = 1, \dots, 183$). We focus on weak-form efficiency, i.e., the information set used for the forecasts consists of past spot and futures rates. Thus, the vector $X_{i,t}$ in equation (2) includes the variables r_{t-1} , $f_{i,t-1}$ and the forecast innovation $f_{i,t} - f_{i-1,t}$.⁷

Detailed results are presented in Tables A2 to A7 in the Appendix. They are summarized in Table 1.⁸ In all regressions, the constant α turns out to be insignificant and the β coefficient is not significantly different from 1. Accordingly, Euribor futures rates are unbiased predictors of future spot rates. Futures rates with forecast horizons of less than four months fulfil the conditions for informational efficiency, i.e. the coefficients on past spot and futures rates are insignificant and the error terms show no serial correlation. In contrast, for futures rates with longer forecast horizons, past spot and futures rates are insignificant, but the error process exhibits significant first order serial correlation.⁹ Thus, in these cases, we accept unbiasedness but reject efficiency.

⁶Figure A1 in the Appendix shows the root mean squared error in dependence of the days to maturity, and confirms this presumption. The increasing graph shows that the predictive accuracy diminishes as the contract horizon is extended.

⁷Since the forecast errors of $f_{i-1,t}$ and $f_{i,t}$ overlap, the only new information for the investor between these two days is the innovation or the difference between these two futures rates.

⁸A detailed description of the test we performed to investigate the structure of the error covariance matrix is listed in Table A1 in the Appendix.

⁹Since our sample extends over the millennium change, we also tested for an increase in the risk premium for the March 2000 contract by including a Y2K dummy in our regressions. This dummy takes the value of one for market days between 13 September 1999 and 7 January 2000, and zero otherwise. The Y2K dummy turns out to be insignificant in all regressions. So, the considerably increased intraday volatility and transaction costs documented by Hartmann et al. (2001) for the euro overnight market during the Y2K changeover week did not seem to impair, in any significant fashion, the ability of the money market to predict short-term rates.

Table 1: Summary of Results of Efficiency Tests with Overlapping Data

Forecast Horizon	Information Set $X_{i,t}$	Unbiasedness $\alpha = \beta - 1 = 0$	Efficiency	
			$\alpha = \beta - 1 = \gamma = 0$	$\rho = 0$
1–31 days	–	A ¹⁾	A	a ²⁾
	r_{t-1}	A	A	a
	$f_{i,t-1}$	A	A	a
	$f_{i,t} - f_{i-1,t}$	A	A	a
1–61 days	–	A	A	a
	r_{t-1}	A	A	a
	$f_{i,t-1}$	A	A	a
	$f_{i,t} - f_{i-1,t}$	A	A	a
1–91 days	–	A	A	a
	r_{t-1}	A	A	a
	$f_{i,t-1}$	A	A	a
	$f_{i,t} - f_{i-1,t}$	A	A	a
1–122 days	–	A	A	a
	r_{t-1}	A	A	a
	$f_{i,t-1}$	A	A	a
	$f_{i,t} - f_{i-1,t}$	A	A	a
1–153 days	–	A	A	R ³⁾
	r_{t-1}	A	A	R
	$f_{i,t-1}$	A	A	R
	$f_{i,t} - f_{i-1,t}$	A	A	R
1–183 days	–	A	A	R
	r_{t-1}	A	A	R
	$f_{i,t-1}$	A	A	R
	$f_{i,t} - f_{i-1,t}$	A	A	R

Notes: ¹⁾F-Test accepted at a significance level of 10%. ²⁾LM-Test accepted at a significance level of 10%. ³⁾LM-Test rejected at a significance level of 5%.

To summarize, the results indicate that financial markets were, on average, able to predict Euribor futures rates during the first years of EMU.

III. The Impact of Monetary Policy Announcements

A. Volatility of Euribor Futures Rates

As pointed out by Poole and Rasche (2000), the EMH implies that futures rates should react to the announcement of monetary policy decisions, if and only if these announcements contain some news about future money market rates. This suggests that one can use the effect of such announcements on the volatility of futures rates as a measure for the quality of the central bank's

information management. If the volatility of futures rates is systematically larger on days when monetary policy decisions are announced than otherwise, then these announcements typically carry significant news or lead to a new interpretation of the central bank's intentions by the markets. Testing for announcement effects on volatility is, therefore, a check on the quality of the central bank's information management.

During our sample period, 92 meetings of the ECB Governing Council occurred, including 16 at which central bank interest rates were changed; see Table 2. The ECB Governing Council usually met on the first and the third Thursday of a month. Only six Council meetings took place on Wednesdays,¹⁰ and one, the meeting of 17 September 2001, on a Monday. In November 2001 the Governing Council announced that – as a rule – it would assess its monetary policy stance only in the first meeting of the month. Accordingly, since then it has not changed interest rates on any of the second meetings during any month.

Table 2: ECB Interest Rate Changes between November 1999 and September 2003

Decisions on	Deposit rate (%)	MRO rate (%)	Marg. lending rate (%)
21 Jan 99	2.00	3.00	4.00
8 Apr 99	1.50	2.50	3.50
4 Nov 99	2.00	3.00	4.00
3 Feb 00	2.25	3.25	4.25
16 Mar 00	2.50	3.50	4.50
27 Apr 00	2.75	3.75	4.75
8 Jun 00	3.25	4.25	5.25
31 Aug 00	3.50	4.50	5.50
5 Oct 00	3.75	4.75	5.75
10 May 01	3.50	4.50	5.50
30 Aug 01	3.25	4.25	5.25
17 Sep 01	2.75	3.75	4.75
8 Nov 01	2.25	3.25	4.25
5 Dec 02	1.75	2.75	3.75
6 Mar 02	1.50	2.50	3.50
5 Jun 03	1.00	2.00	3.00

Notes: MRO = main refinancing operation

The interest rate change at the 4 January 1999 meeting is not included in this table, because it was already decided in December 1998

Source: ECB

¹⁰They were the meetings on 2 June 1999, 15 December 1999, 5 January 2000, 21 June 2000, 11 April 2001 and 23 May 2001.

We measure volatility in terms of the absolute change in the futures rate between two trading days multiplied by 100, $\sigma_t = 100 |f_{i,t} - f_{i-1,t}|$. To avoid double counting of ECB Council meetings, we restrict the analysis to the closing rates of all nearby three-month Euribor futures between 1 January 1999 and 15 September 2003. Thus, we consider futures rates with forecast horizons of one to 91 days.

As a first step, we regress the volatility on a constant and four week-day dummies to check for any week-day effect that might be caused by the money market's microstructure. As before, we employ a panel estimator correcting for serial correlation and heterogeneity across groups.¹¹ The results are shown in Table 3. They indicate a significant Thursday effect, in line with earlier findings by Hartmann et al. (2001). In the first half of the sample, this Thursday effect was more than twice the size estimated in the second half of the sample period.¹²

Figure 7 depicts the volatility of futures rates on Governing Council days. Meetings at which the ECB Council changed its policy rates are highlighted. On 38 out of 92 Governing Council days (41% of all Council meetings), the announcement effect was small, causing a volatility of less than 1.72 basis points, which is the average volatility on all days plus two standard deviations of 0.10 basis points (compare Table 3). These policy decisions apparently incorporated only little new information. On 29 meetings (32% of all Council meetings), the volatility exceeded 3.04 basis points, twice the average volatility on all days.

Table 3: Day-Average Volatilities with Corrections for Heteroskedasticity

	$ f_{i,t} - f_{i-1,t} $		
	March 99–Sep 03	March 99–June 01	Sep 01–Sep 03
const.	1.52 [0.00]	1.47 [0.00]	1.57 [0.00]
Tuesday	0.12 [0.51]	0.22 [0.37]	0.02 [0.95]
Wednesday	0.26 [0.17]	0.46 [0.13]	0.03 [0.87]
Thursday	0.81 [0.00]	1.12 [0.02]	0.47 [0.02]
Friday	0.24 [0.12]	0.37 [0.10]	0.11 [0.50]
R ²	0.02	0.02	0.01
Obs.	1185	615	570

Notes: P-Values are reported in squared brackets. Volatilities are multiplied by 100 to express basis points.

¹¹Group-wise heterogeneity is suggested by the 'Samuelson effect' which holds that the volatility of asset prices increases with decreasing maturity i ; see Samuelson (1965).

¹²A coefficient test rejects at a 5% significance level the null-hypothesis that the Thursday effect is the same in both sub-samples.

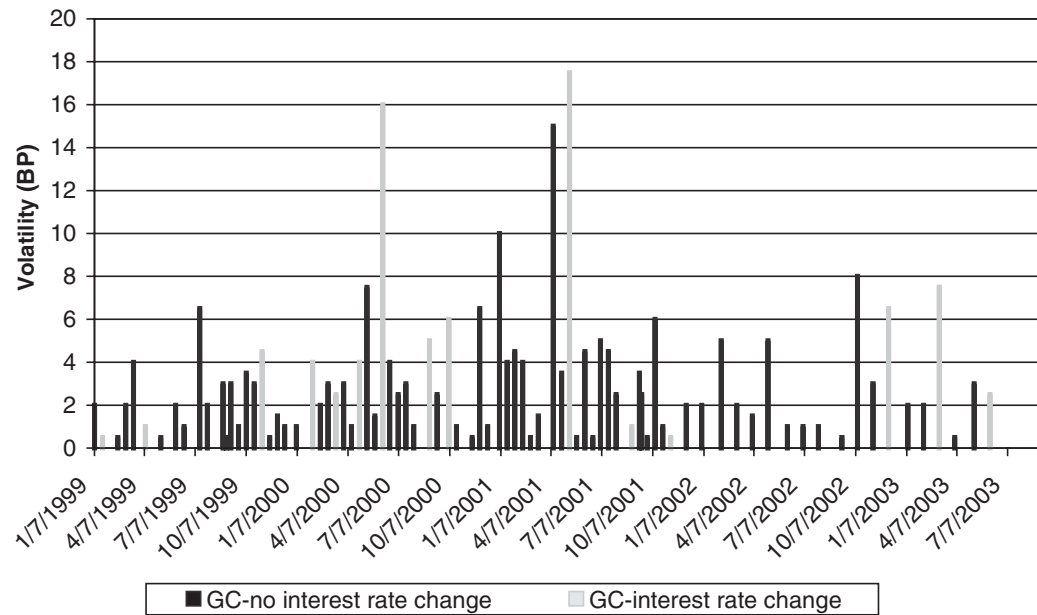


Figure 7: Volatility of Futures Rates at Governing Council Meetings

Figure 7 shows that there are only four days with a volatility of more than 10 basis points: 8 June 2000, 4 January 2001, 11 April 2001 and 10 May 2001. Half of these are days on which the ECB announced an interest rate change. Obviously, markets were to some extent surprised by these policy decisions. We do not observe similarly large volatilities after May 2001, indicating that the ECB's communication policy managed to avoid large surprises in the second half of the sample.

On 11 May 2000, the Governing Council decided not to change its monetary policy, but nevertheless Euribor futures rates decreased by almost 8 basis points. A month later, on 8 June 2000, the Governing Council announced a rate cut of 50 basis points, and Euribor futures rates decreased by another 16 basis points. Thus, market participants apparently expected correctly an interest rate cut, but they were surprised by the size of the policy change. The high volatility on 4 January 2001 is probably the reaction to the Federal Reserve funds rate cut of 50 basis points the day before. Financial markets expected the ECB to follow the Federal Reserve's lead and, therefore, the Euribor futures rate fell by 10 basis points. On 11 April 2001, the ECB announced it would hold interest rates constant and futures rates reacted with a rate jump upwards of 15 basis points. At the Council meeting one month later, the ECB decided to cut interest rates by 25 basis points, and financial markets reacted with a downward correction of futures rates by more than 18 basis points. This volatility pattern indicates that market participants correctly expected the direction of interest rate changes intended by the ECB, but there was uncertainty about the timing. Thus,

the markets misinterpreted the ECB's inaction at the first Council meeting as a signal that rates would not be changed for a while, hence they were surprised by the subsequent move.

The upper panel of Table 4 presents the results of regressing the volatility of Euribor futures rates on a constant and three dummies: the Thursday dummy, one dummy for all Council meetings and one for Council meetings with announced interest rate changes. The Thursday dummy is now insignificant. This suggests that the apparent Thursday effect in Table 3 relates to the Governing Council meetings.

In the lower panel of Table 4, we drop the Thursday dummy. Between 1999 and 2003 the volatility of Euribor futures on Governing Council meeting days exceeded the volatility on other days by 0.8 basis points. This effect is significant at the 10% level over the entire sample. On Governing Council days with announced monetary policy changes, Euribor futures rates are the most volatile. During the first two years of EMU, the extra volatility was 3.93 basis points and significant at the 1% level. In contrast, between September 2001 and September 2003, the extra volatility was only one basis point and it was not statistically significant anymore.

The empirical results thus indicate that, during the first two years of EMU, some decisions of the ECB Council were not fully anticipated by market participants. Even in the first two years, however, the impact of policy announcements on the volatility of futures rates was small compared to the typical size of changes in central bank interest rates, which is 25 or 50 basis points. This suggests that the economic importance of the announcement effect was limited. With our new methodology based on euro money market

Table 4: Effects of Monetary Policy Decisions on Volatility

	$ f_{i,t} - f_{i-1,t} $		
	March 99–Sep 03	March 99–June 01	Sep 01–Sep 03
const.	1.67 [0.00]	1.73 [0.00]	1.61 [0.00]
Thursday	0.31 [0.16]	0.38 [0.36]	0.24 [0.20]
GC Days	0.56 [0.08]	0.50 [0.33]	0.57 [0.19]
GC with Int. Rate Change	2.74 [0.00]	3.88 [0.00]	1.03 [0.20]
R ²	0.04	0.06	0.02
const.	1.72 [0.00]	1.79 [0.00]	1.65 [0.00]
GC Days	0.80 [0.08]	0.78 [0.06]	0.78 [0.07]
GC with Int. Rate Change	2.75 [0.00]	3.93 [0.00]	0.99 [0.22]
R ²	0.04	0.06	0.02
Obs.	1185	615	570

Notes: P-Values are reported in squared brackets. Volatilities are multiplied by 100 to express basis points.

futures, we therefore confirm some of the general results of earlier papers referred to in the introduction.

B. Announcement Effects on Prediction Errors

A decline in the absolute prediction error between two trading days, $100(|r_t - f_{i,t}| - |r_t - f_{i-1,t}|)$, indicates that the predictability of future spot rates improved due to market developments on that day. In connection with announcements of ECB monetary policy decisions,¹³ this would suggest that the announcement improved the markets' ability to predict future spot rates.

Figure 8 shows the changes of the absolute prediction error on Governing Council and non-Council days. There are six outliers, i.e., changes of at least ten basis points in absolute values. Four of these signal a large improvement in interest rate predictability. Three of the six outliers are observed on Governing Council days, two of them (the Council meetings on 11 April 2001 and on 10 May 2001) show a decrease and one (the Council meeting on 4

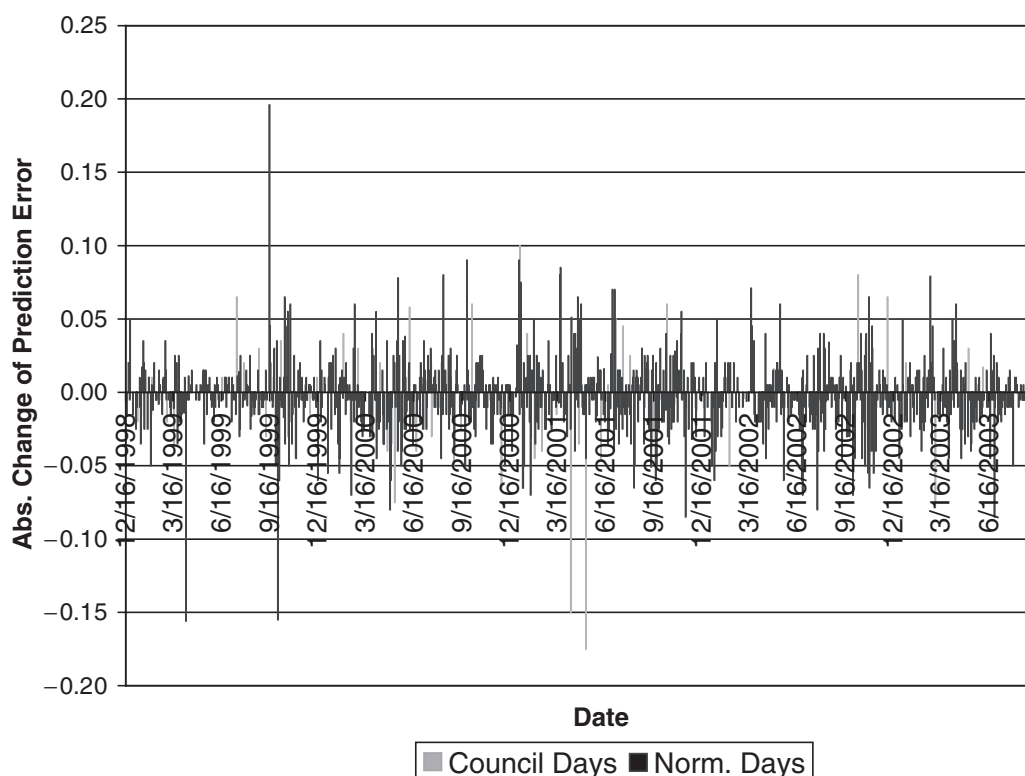


Figure 8: Total Change of the Prediction Error of Euribor Futures

¹³We also regressed the absolute change of the futures rates' prediction error on weekday dummies. The results show no systematic effects. The estimation results are available from the authors on request.

January 2001) an increase in the prediction error. In general, positive and negative changes occur with similar frequencies. Thus, the graphical analysis suggests that the information released on Governing Council meetings did not systematically improve the markets' ability to forecast interest rates.

A regression of the changes in absolute prediction errors on a constant and dummies for Council days and Council days with interest rate changes confirms that impression. We use again an OLS estimation approach with panel-corrected standard errors to capture cross-sectional heteroskedasticity and correlation. Table 5 shows that the coefficients on both dummies are statistically not significant. ECB policy announcements did not systematically help market participants to improve their interest rate forecasts.

Table 5: Effects of Monetary Policy Announcements on Day-Prediction Error

	$ r_t - f_{i,t} - r_t - f_{i-1,t} $		
	March 99–Sep 03	March 99–June 01	Sep 01–Sep 03
const.	– 0.26 [0.00]	– 0.18 [0.02]	– 0.34 [0.00]
GC Days	– 0.01 [0.98]	– 0.57 [0.28]	0.90 [0.13]
GC with Int. Rate Change	– 0.99 [0.29]	– 0.98 [0.48]	– 1.11 [0.38]
R ²	0.04	0.06	0.02
Obs.	1185	615	570

Notes: P-Values are reported in squared brackets. Volatilities are multiplied by 100 to express basis points.

IV. Conclusion

In this paper, we analyse three aspects of the predictability of money market rates in the EMU. The first is the efficiency of the three-month Euribor interest rate futures markets. The estimation results show that Euribor futures rates with a forecast horizon of up to four months are unbiased and informationally efficient predictors of future spot rates.

The second aspect is the impact of the ECB's monetary policy decisions on the volatility of the Euribor futures rates. The efficient market hypothesis (EMH) implies that futures rates only change between two days when new information comes on the market. The volatility of the futures rates at Governing Council meetings can therefore be used as a measure of surprise caused by the central bank's policy decision. Estimation results show that,

during the first five years of EMU, the average volatility of the Euribor futures rates on Governing Council days was significantly larger than on non-Council days, and most of that extra volatility came from Governing Council meetings at which interest rate changes were adopted. During the first two years, the volatility of futures rates after policy actions was about three times as large as during the subsequent two years of EMU. A closer look at the futures rate changes at Governing Council days shows that a majority of ECB policy decisions were anticipated correctly by the markets, and only a few constituted an economically significant surprise. We do not explore in greater depth, however, what could have led European central bankers to surprise markets at those instances. This could be usefully done in future research.

The third aspect is the impact of Governing Council meetings on the change in the absolute prediction error of the Euribor futures rates. Our analysis suggests that the information released on Governing Council days did not systematically improve the markets' ability to forecast interest rates.

The predictability of short-term money market rates is an indicator of monetary policy transparency and the effectiveness of the central bank's information policy. Our empirical results suggest that participants in the new euro money markets were able to predict short-term rates well. Moreover, the predictability of ECB policy decisions seems to have further improved during the first years of EMU. ECB Council decisions still cause some surprises, but their effect on volatility is small. In sum, the ECB's information management in this new environment has been met with considerable success.

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Appendix

Table A1 shows the test results of the structure of the error term matrix. A Chow test does not reject the non-existence of individual effects. This shows that risk premia do not vary with the length of the forecast horizon i of futures rates, $\alpha_i = \alpha_j$ for $i \neq j$.

A Lagrange Multiplier (LM) test rejects the null hypothesis of a common variance across panels at every reasonable significance level. Accordingly, the variance of the error process differs depending on the days to maturity i .

A White test rejects the null hypothesis of homoskedasticity across time t for all panel regressions and futures rates with different forecasting horizons.

The results of a Breusch Pagan LM test show that we can reject, at every significance level, the null-hypothesis of no cross-sectional correlation. That means that if a futures rate has a positive prediction error i days before contract maturity, it is likely that the futures rate will show a similar prediction error j days before maturity, where $j = [i - n, i + n]$ and n denotes a sufficiently small time distance to this specific day.

The LM test results show that we cannot reject the null hypothesis of no serial correlation in the regressions with futures rates with a forecast horizon of up to four months.

Table A1: Test Results

Forecast Horizon	$X_{i,t}$	Indiv. Effects $H_0: \alpha_i = \alpha$	Heterosk. across i $H_0: \sigma_{ij} = \sigma$	Contemp. Correl. $H_0: \sigma_{ij} = 0$	Serial Correl. $H_0: \rho = 0$
1-31 days	–	0.92	0.00	0.00	0.49
	r_{t-1}	0.76	0.00	0.00	0.36
	$f_{i,t-1}$	0.77	0.00	0.00	0.63
	$(f_{i,t} - f_{i-1,t})$	0.92	0.00	0.00	0.50
1-61 days	–	0.97	0.00	0.00	0.52
	r_{t-1}	0.89	0.00	0.00	0.95
	$f_{i,t-1}$	0.82	0.00	0.00	0.63
	$(f_{i,t} - f_{i-1,t})$	0.95	0.00	0.00	0.31
1-91 days	–	1.00	0.00	0.00	0.69
	r_{t-1}	1.00	0.00	0.00	0.89
	$f_{i,t-1}$	1.00	0.00	0.00	0.74
	$(f_{i,t} - f_{i-1,t})$	1.00	0.00	0.00	0.82
1-122 days	–	0.93	0.00	0.00	0.16
	r_{t-1}	0.92	0.00	0.00	0.32
	$f_{i,t-1}$	1.00	0.00	0.00	0.42
	$(f_{i,t} - f_{i-1,t})$	0.92	0.00	0.00	0.20
1-153 days	–	0.59	0.00	0.00	0.00
	r_{t-1}	0.65	0.00	0.00	0.00
	$f_{i,t-1}$	1.00	0.00	0.00	0.01
	$(f_{i,t} - f_{i-1,t})$	0.57	0.00	0.00	0.00
1-183 days	–	0.91	0.00	0.00	0.00
	r_{t-1}	0.99	0.00	0.00	0.00
	$f_{i,t-1}$	1.00	0.00	0.00	0.00
	$(f_{i,t} - f_{i-1,t})$	0.90	0.00	0.00	0.00

Figures in P -values

Table A2: Estimation Results for Euribor Futures Rates with 1-Month Forecast Horizon

Equ.	Independent Variables					Statistics				Coefficient test			
	α	β	r_{t-1}	$f_{i,t-1}$	$f_{i,t} - f_{i-1,t}$	R^2	RMSE	NT	AR1	α	β	$\gamma_{1,2,3}$	F -test
(3)	0.07	0.98				1.00		393	0.49	0.09	0.06		0.15
	0.09	0.00											
(4)	0.06	0.97	0.01			1.00		372	0.36	0.17	0.22	0.66	0.27
	0.17	0.00	0.66										
(5)	0.07	0.97		0.01		1.00		366	0.63	0.14	0.26	0.76	0.29
	0.14	0.00		0.76									
(6)	0.07	0.98			0.15	1.00		389	0.50	0.10	0.07	0.22	0.16
	0.10	0.00			0.22								

Notes: Standard errors in parentheses; R^2 is the proportion of the total variation in $r_{i,t}$ explained by the regression; RMSE denotes the root mean squared error; NT is the number of observations; AR1 gives the P -values of a LM test for first-order autocorrelation; the individual coefficient tests are simple linear tests of coefficient restrictions against the null hypothesis where $\alpha = 0$, $\beta = 1$, $\gamma_{1,2,3} = 0$; test results in P -values.

Table A3: Estimation Results for Euribor Futures Rates with 2-Month Forecast Horizon

Equ.	Independent Variables					Statistics				Coefficient test			
	α	β	r_{t-1}	$f_{i,t-1}$	$f_{i,t}-f_{i-1,t}$	R^2	RMSE	NT	AR1	α	β	$\gamma_{1,2,3}$	F-test
(3)	0.05 0.46	0.98 0.00				0.99	0.10	806	0.52	0.45	0.28		0.27
(4)	0.03 0.72	0.95 0.00	0.03 0.51			0.99	0.10	763	0.95	0.71	0.29	0.51	0.40
(5)	0.03 0.69	0.96 0.00		0.02 0.58		0.99	0.10	752	0.63	0.69	0.34	0.58	0.43
(6)	0.05 0.46	0.98 0.00			0.06 0.68	0.99	0.10	798	0.31	0.46	0.28	0.68	0.39

See notes for Table A2

Table A4: Estimation Results for Euribor Futures Rates with 3-Month Forecast Horizon

Equ.	Independent Variables					Statistics				Coefficient test				
	α	β	r_{t-1}	$f_{i,t-1}$	$f_{i,t}-f_{i-1,t}$	Y2K	R^2	RMSE	NT	AR1	α	β	$\gamma_{1,2,3}$	F-test
(3)	0.02	0.99				0.19	0.98	0.13	1204	0.69	0.63	0.82		0.57
	0.82	0.00				0.27								
(4)	0.01	0.98	0.01			0.19	0.98	0.13	1143	0.89	0.89	0.71	0.85	0.78
	0.89	0	0.85			0.27								
(5)	0.03	0.99		0.00		0.19	0.98	0.13	1114	0.74	0.77	0.84	0.95	0.74
	0.77	0.00		0.95		0.28								
(6)	0.02	0.99			-0.01	0.20	0.98	0.13	1183	0.82	0.82	0.63	0.95	0.74
	0.82	0.00			0.95	0.26								

See notes for Table A2

Table A5: Estimation Results for Euribor Futures Rates with 4-Month Forecast Horizon

Independent Variables														Statistics				Coefficient test			
Equ.	α	β	r_{t-1}	$f_{i,t-1}$	$f_{i,t}-f_{i-1,t}$	Y2K	R^2	RMSE	NT	AR1	α	β	$\gamma_{1,2,3}$	F-test							
(3)	0.05	0.98				0.25	0.96	0.17	1579	0.16	0.67	0.43		0.29							
	0.67	0.00				0.24															
(4)	0.05	0.98	-0.01			0.25	0.96	0.18	1518	0.32	0.69	0.80	0.92	0.48							
	0.69	0.00	0.92			0.24															
(5)	0.10	1.01		-0.04		0.23	0.96	0.17	1465	0.42	0.40	0.89	0.49	0.39							
	0.40	0.00		0.49		0.27															
(6)	0.05	0.98			-0.05	0.25	0.96	0.17	1555	0.20	0.67	0.43	0.75	0.45							
	0.67	0.00			0.75	0.24															

See notes for Table A2

Table A6: Estimation Results for Euribor Futures Rates with 5-Month Forecast Horizon

Equ.	Independent Variables				Statistics			Coefficient test					
	α	β	$r_{t-1}f_{i,t-1}$	$f_{i,t}-f_{i-1,t}$	Y2K	R ²	RMSE NT	AR1	α	β	$\gamma_{1,2,3}$	F-test	
(3)	0.12	0.95			0.12	0.92	0.21	1988	0.00	0.45	0.27	0.26	
	0.45	0.00			0.62								
(4)	0.10	0.92	0.03		0.12	0.92	0.22	1927	0.00	0.58	0.40	0.74	0.45
	0.58	0.00	0.74		0.62								
(5)	0.16	0.98	− 0.03		0.14	0.93	0.21	1846	0.01	0.36	0.77	0.7	0.42
	0.36	0.00	0.70		0.56								
(6)	0.12	0.95		− 0.11	0.12	0.92	0.21	1960	0.00	0.45	0.27	0.51	0.39
	0.45	0.00		0.51	0.62								

See notes for Table A2

Table A7: Estimation Results for Euribor Futures Rates with 6-Month Forecast Horizon

Equ.	Independent Variables				Statistics			Coefficient test					
	α	β	$r_{t-1}f_{i,t-1}$	$f_{i,t}-f_{i-1,t}$	Y2K	R ²	RMSE NT	AR1	α	β	$\gamma_{1,2,3}$	F-test	
(3)	0.15	0.94			0.14	0.88	0.25	2368	0.00	0.46	0.28	0.26	
	0.46	0.00			0.59								
(4)	0.10	0.82	0.13		0.12	0.88	0.25	2307	0.00	0.68	0.12	0.31	0.29
	0.68	0.00	0.31		0.67								
(5)	0.18	0.95	− 0.02		0.16	0.89	0.25	2193	0.00	0.41	0.61	0.86	0.45
	0.41	0.00	0.86		0.54								
(6)	0.15	0.94		− 0.10	0.15	0.89	0.25	2313	0.00	0.46	0.27	0.57	0.41
	0.46	0.00		0.57	0.57								

See notes for Table A2

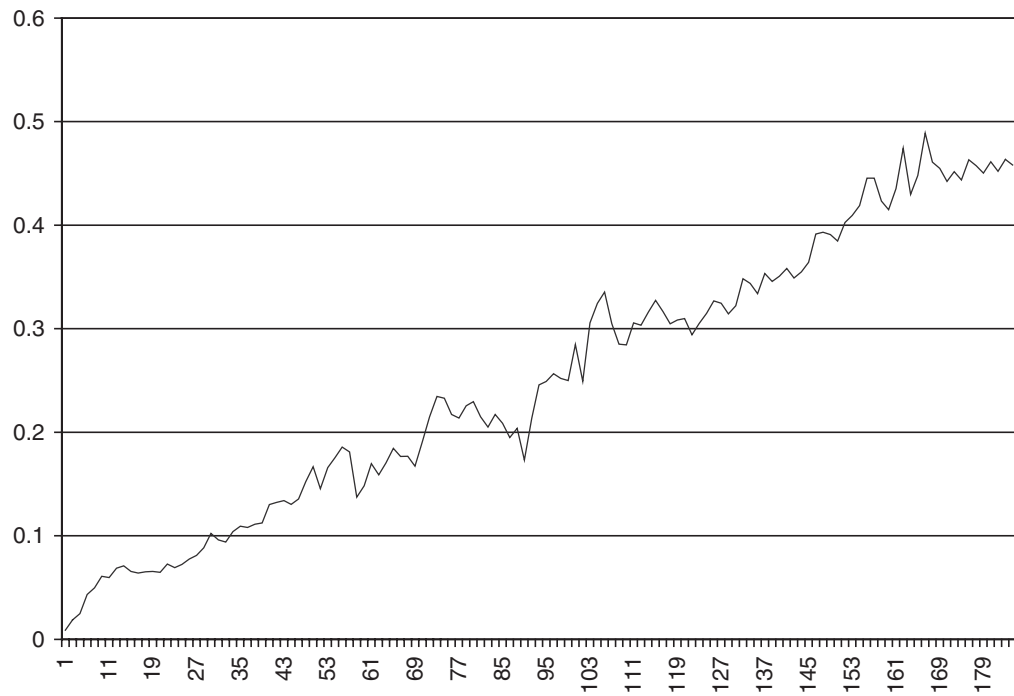


Figure A1: The Root Mean Squared Error of Euribor Futures from Spot Rates

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