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ECONOMETRICS | RESEARCH ARTICLE

The dynamic of bank stock price and its fundamentals: Evidence from Indonesia

Hendrik Widjaja¹* and Moch. Doddy Ariefianto¹

Abstract: This study examines the relationship between bank stock price and its selected fundamentals, namely, profitability, credit risk, and liquidity risk. Using the dynamic common correlated effect (DCCE) technique, we discover a mechanism error-correction between the stock price and the selected fundamentals. We estimate that the equilibrating process of stock price takes between 2.62 and 3.22 months. This study also provides significant support for hypotheses of the positive role of profitability (proxied by ROE and NIM) to bank stock price. Credit and liquidity risk measures do not significantly affect stock price.

Subjects: Environment & Business; Environment & Economics; Econometrics; Banking

Keywords: Stock price; profitability risk; credit risk; liquidity risk

EL Classification: C23; G3; G18; G211



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financial technology. He has published his works in various reputable international journals.

1. Introduction

Banks are inherently hard to value (more opaque) than companies in other industries (Blau et al., 2017; Morgan, 2002). Existing empirical studies reveals the difficulty in assessing the value of banks by outsiders (Jones et al., 2012). This difficulty is arised from asymmetric information and moral hazard problem that is inherent in a bank business model (Greenbaum et al., 2019, page 8). These problems have caused substantial measurement error in the loan and scurities, which are the dominant part of banks asset (Freixas & Rochet, 2008 page 127).

Banks heavily apply risk management techniques to reduce the impact of asymmetric information and moral hazard to asset value (therefore, bank is often called a business of handling risk, Koch & MacDonald, 2014 page 22). Nevertheless, its efficacy is still fell short of expectation, which explains heavy handling of this industry by regulators (Barth et al., 2008, page 12). Due to previously explained narrative, market discipline (by investor) is also believed to be less effective for banks (Morgan & Stiroh, 2001).

There are several common profitability measures alternatives for banks like Earnings per Share (EPS), Return on Asset (RoA), Earnings Growth, and Return on Equity (ROE). In this study, we use ROE; as doccumented by Pennacchi and Santos (2021), this measure is the most common form of profitability measurement, which is often used in setting management compensation contracts and investor communications.

As a business in handling risk, it is only natural to expect the response of its stock investors highly dependent not only on the level of profit achievement but more importantly on how is it achieved? Does the profitability bring hidden issues that could materially reduce or even subsequently reverse the achievement? The market power and pricing behavior of banks depend on the broader size of the bank's margins beyond interest income, and because of the complexity of the bank's business, any policy analysis of a bank's market power must rely on different indicators and especially on indicators that reflect the full reality of the bank's output mix and explicitly considers the prices of all outputs relative to their marginal costs (Valverde & Fernández, 2007). Investors face a high-dimensional forecasting problem; this are behavioral biases, and investors arguably face a high-dimensional prediction problem in the real world. Investor heterogeneity can result in the role of endogenous price-based signals from which investors can extract information about not only asset fundamentals but also the trading behavior of other investors and need to learn (Martin & Nagel, 2022). In this regard, it is standard practice by bank stock investors to view financial performance in conjunction with accompanying risks

Bank business is subject to various kinds of risks, but there are two most prominent risks that are relevant for Indonesia context, liquidity risk and credit risk (Saunders & Cornett, 2004 page 173).¹ There are ample empirical studies that support this notion. Credit risk, liquidity risk, and bank capital were shown to affect bank profitability in either a positive or negative way. Banks need to change their credit policies, which aim to reduce credit risk, whereas good credit policies lead to reduced bad credit in banks and thus improved profitability. In addition, the banks should have more liquidity and higher capital in order to face any future situations that might have an effect on their profitability (Saleh & Abu Afifa, 2020). Credit risk and liquidity risk on determining banks' cost of intermediation and their effects on bank margins are sensitive to economic fluctuations (Kesraoui et al., 2022). The impact of liquidity on profitability is positively correlated, while the credit risk negatively and significantly impacts the profitability (Abbas et al., 2019).

Information in a stock price is divided into three components, namely, past information, current information available to the public, and current private information. The extent and amount of this information reflected in stock price reflect market efficiency. Blau et al. (2017) argued that since risk in the bank intermediation process is quite high, the information carried in its stock price is limited. Investors can only observe efforts of risk management, but not its efficacy, and hence, their decision is based on less then perfect information resulting in error. Investors learn from

experience and therefore act more in accordance with the prescriptions of finance theory as long as outcomes are observable and salient. They view financial education, decision support, default options, salient disclosure, and private investments in financial literacy via investment mistakes as complements, not substitutes. Koestner et al. (2017); Carlin and Robinson (2012). Surely, investors learn from their mistakes and revise their decision accordingly.

Previous studies in the relationship of bank stock return and its determinants mostly cast in static design without accounting for feedback and learning effects. We contribute to the existing literature by a study conducted in a dynamic paradigm. Our dynamic empirical design allow us to not only learn about long-run relationship but also uncover the error-correction mechanism (ECM): a manifestation of feedback and learning process. To the best of our knowledge, we are the first to perform this kind of study.

To meet our research objective, we construct empirical design using listed bank stock price as a dependent variable and profitability, credit risk, and liquidity risk as our variables of interest and macroeconomic variables (interest rate, exchange rate change, and GDP growth) as control. We apply a novel econometric technique, namely, Dynamic Common Correlated Effect (DCCE) (Chudik & Pesaran, 2015), to a monthly long panel data set that comprises 21 banks the for period of July 2006 to August 2019. The flow of application of the econometric technique follows closely Ariefianto et al. (2021).

Previous studies proposed that lack of transparency can reduce stock price stability and lead to a lower level of market efficiency (Blau et al., 2017; Haggard et al., 2008; Jin & Myers, 2006). Due to a lack of company-specific information, Veldkamp (2006) created a model in which investors must rely on broad signals to estimate firm cash flows. Moreover, lack of information leads to greater heterogeneity among securities, and this results in uninformed share prices.

When two or more prices of security allow an investor to build a net investment portfolio that will produce a definite return, an arbitrage opportunity (risk-free) arises. These opportunities initiate a large trading volume that puts pressure on the security price. According to Bodie et al. (2018; 326), this pressure will continue until the price reaches the level that prevents arbitrage.

Bank opacity may affect the ability of outsiders to precisely evaluate the bank, and hence, bank stock price information may also be inaccurate. According to Simoens and Vander Vennet (2021), profitability was one of the primary drivers of bank M/B ratios from 2007 to 2017. This changed after the crises as improved cost efficiency was a major factor in increased profitability. Furthermore, an extended period of low interest rates in Europe could have a detrimental impact on M/B valuations and bank franchise values. This analysis suggests that higher deposit shares were related to lower performance in post-crisis Europe and could be attributed to the low-for-long monetary policy environment reducing the value of bank funding.

According to Yuksel and Zengin (2017), the net interest margin is an estimate of the diference between the weighted average yield on interest income and expenses (Islam & Nishiyama, 2016). Bank profitability is usually measured by net interest margin because it is one of the useful indicators to describe the bank's core income capability where net interest margin focuses on measuring profits from interest activities (Batten & Vo, 2019). Our hypothesis that higher profitability positively correlates with the stock price. For example, assuming that the net interest margin is not good will affect the net income and the bank's stock price.

Hypothesis 1. Bank stock price positively correlates with profitability

The profitability variables adopted for the study were NIM and ROE. NIM measures the gap between what the bank pays savers and what the bank receives from borrowers. Thus, NIM focuses on the traditional borrowing and lending operations of the bank (Chen et al., 2018; Gadzo et al., 2014). NIM is also the difference between interest income generated (by banks) from loans and interest expense paid out to depositors, relative to average earning assets. It shows the profit a bank is making from its basic function, out of its average earning assets. Higher NIM convinces banks to give out more loans and hence increase its credit and operational risk management levels. Studies that have employed NIM as a profitability measure include Heffernan and Fu (2008), Nguyen (2012), Lee et al. (2014), Noman et al. (2015), and Gadzo et al. (2019).

The bank would surely select ROE assuming that they want to create a rosy picture of its performance. The use of ROE in bank compensation contracts and investor communications is consistent with this decision (Pennacchi & Santos, 2021). Based on this, ROE is estimated to positively correlate with stock prices.

Hypothesis 2. Bank stock price negatively correlates with credit risk

Imbierowicz and Rauch (2014) suggests that liquidity and credit risk jointly affect bank health in a sample of US commercial banks. According to the findings of Vazquez and Federico (2015), simultaneous exposure to liquidity and credit risk increases bank troubles during the crisis based on a collection of European and American banks. Furthermore, liquidity and credit risks are believed to be regulated together (Acharya, Mehran and Thakor, 2016; Acharya & Mora, 2015; Brunnermeier et al., 2009; Calomiris et al., 2015; Distinguin et al., 2013; He & Xiong, 2012; Imbierowicz & Rauch, 2014; Vazquez & Federico, 2015). However, since banks are heavily reliant on the interbank market, increasing capital requirements might be seen as a prudential step to reduce liquidity and insolvency concerns.

Hypothesis 3. Bank stock price positively correlates with liquidity measures

According to Phung et al. (2021), non-performing loans negatively correlate with bank efficiency. NPL can be seen as an indicator of a bank's asset health. This is because the bulk of assets that might produce income to a bank are from loans. The liquidity risk of a bank increases mainly as the NPL ratio increases, and this reduces the value of productive assets and can decrease the value of the bank's shares. Once the value of the productive assets decreases significantly higher, NPL can negatively correlate with the stock price.

Higher credit risk is negatively correlated with stock price, while liquidity risk positively is correlated with stock price.

2. Methodology and data

Based on hypothesis contruction explained in the previous section, we construct the relationship of stock price and the explanatory variables in the form of short-run (Error-Correction Model) and long-run regression as follows:

$$\Delta SP_{i,t} = \beta_0 + \beta_1 \Delta Prof_{it} + \beta_2 \Delta Cred_{it} + \beta_3 \Delta Liq_{it} + \sum_{i=4}^{10} \beta_i CV_i + \delta ECT_{t-1} + e_{i,t},$$
(1)

Table 1. Variables, proxies, and expected sign. The research hypotheses are given as expected impact sign against the dependent variable from the correlation matrix, and we have also considered the real facts in the Bank and also used the previous study as a reference

No of Variables		Proxy	Expected sign (hypothesis)
Long-run and short-run equ	uations		
- Dependent variable			
1	Stock price	Share price (SP)	
- Independent variable			
2	Profitability (Prof)	Net interest margin (NIM)	Positive
		Operating income/equity (<i>ROE</i>)	Positive
3	Credit risk (Cred)	Non-performing loan/ loan (NPL)	Negative
		Loan allowance for impairment/loan (LAW)	Negative
4	Liquidity risk (Liq)	Deposit/loan (DLR)	Positive
		Liquid asset/total assets (LAR)	Positive
5	Error correction term	Lagged one of long-run equation residual (ECT)	Negative
Short-run equations/Contro	l variable		
6	SIZE	Log of bank asset (SIZE)	
7	CAPITAL	Equity/total assets (CAPI)	
8	Financial system stability	Exchange rate: IDR per USD (<i>EXCH</i>)	
		Interest rate (INTR)	
9	Structural factor	Gross domestic product riil (GDPR)	

 $ECT_{t-1} = SP_{t-1} - \alpha_1 Prof_{it-1} + \alpha_2 Cred_{it-1} + \alpha_3 Liq_{it-1},$ (2)

where SP is that stock price, Prof is the profitability measure, Cred is the Credit risk measure, Liq is the liquidity measure, and CV is a set of control variables (Size, Capital ratio/capi, Exchange Rate/ EXCH, Interest Rate/INTR, and Riil Gross Domestic Product/GDPR). Equation 1 is the short-run equation, and equation 2 is the long-run equation. Equation 1 and 2 are connected by the errorcorrection term (ECT); it is also a lag one period long-run regression residual whose parameter δ measures the pace of adjustment. This parameter should be negative with the absolute value strictly less than 1. A detailed description on variables, proxies, and their expected sign hypothesis (in relation to the dependent variable) is provided in Table 1.

Our data set comprises monthly series of 21 public banks from July 2006 to August 2019, 3,318 bank monthly observations. The data are collected from the Indonesia Financial Services Authority-OJK (for bank level data), Yahoo finance for the stock price, and Bank Indonesia for exchange rates and interest rates. Appropriate cleansing, interpolation, and winsorization (with cut of 1% and 99%) are applied in data preparation.

Since our data are the type of a long panel (the number of time series is greater than the cross-section), we should pay attention to several important feature: (a) non-stationary (unit root), (b) cointegration, (c) slope heterogeneity, and (d) cross-sectional dependence.

Ideally, our econometric method requires all variables to be included in the model to be nonstationary, i.e. possesses a unit root process. For (bank-level) panel variables, we employ Pesaran (2007) cross-section Augmented Dickey Fuller-PESCADF. This method was developed to be resistant to cross-sectional dependence (Pesaran, 2015b). Unit root tests for time series variables were performed using the ADF approach (Dickey & Fuller, 1979).

Westerlund (2007) developed a method for the cointegration test that takes into consideration the cross-sectional dependence. The rejection of the null hypothesis is interpreted as the presence of cointegration. Persyn and Westerlund (2008) generalize the design to employ a variety of possible dynamic structures. Various cointegration test specifications were performed based on (a) the variable of dependent used (SP); (b) the lag amount used, leads, and LR windows; (c) use or disuse of the automated lag selection (method by Akaike's information criteria; AIC); and (d) the use of the bootstrap approach to obtain robust inference (Blomquist & Westerlund, 2014).

Another issue with data from lengthy panels is the heterogeneity of the cross-section slope rather in addition to the residual component (Eberhardt & Teal, 2011). We test the homogeneity slope against heterogeneity assumption using the method developed by Pesaran and Yamagata (2008) and Blomquist and Westerlund (2013). The test detail covers (a) alternative long-run specifications (SP), (b) the use of standard errors of heteroscedasticity and consistent autocorrelation (HAC), and (c) inclusion of the cross-section mean of the descriptive variables and the lag amount (0–4).

Based on the slope heterogeneity test results, an appropriate version of estimator is selected. For the case of slope heterogeneity, we should use the mean group estimator (Pesaran & Smith, 1995). If the assumption of slope homogeneity is not rejected, then we could use set an identical slope for all cross-section units yielding a pooled mean group estimator reported by Pesaran et al. (1999).

We conduct robustness check to support our empirical finding. The robustness check is performed by (a) running regressions on alternative proxies, (b) using samples based on bank types (state-owned banks: SOE; private-owned: PRIV). and (c) using subsamples based on the period: (i) up to April 2009 (714 observations) and (ii) May 2009—Aug 2019 (2,604 observations). April 2009 was selected as the cut-off period to reflect the condition before and after global financial crisis of 2008.

3. Results and discussion

3.1. Descriptive statistic and preliminary analyses

Table 2 presents the descriptive statistics of the variables used in the study. We can see that the variables are quite well behaved as indicated by closeness of mean and median and no excessive skewness. Banks' profit equals a company's revenues minus expenses. Earning a profit is important to a business because profitability impacts whether a company can secure financing from a creditors, attract investors to fund its operations, and grow its business. Banks must have good governance in managing credit risk and liquidity risk so that banks can generate profits, which will certainly provide satisfaction to shareholders, investors, creditors, and all stakeholders. LAR is a liquid asset ratio, which provides an indication of the liquidity available to meet expected and unexpected demands for cash. The level of liquidity indicates the ability of the deposit-taking sector to withstand shocks to their balance sheet. LAR is calculated by using the core measure of liquid assets as the numerator and total assets as the denominator. LAR is a variable that deserves special attention since it has some degree of skewness. The SP average is 1.72, with a median of 0.60, and the LAR average is 28.38, with a median of 26.08.

Table 3 shows the result of simple correlation (Pearson's Correlation) between variables of interest in this study. The correlation data are represented in triangular form, and multicollinearity does not appear to be a possibility for explanatory variables. The correlation between variables of

: in numeric (up to two digits behind the ic form; and NIM, ROE, NPL, LAW, LDR, ALR	deviation, percentiles, and number of	
used in the study. The variables ard SIZE, EXCH, and GDPR in Logarithm	lan, maximum, minimum, standard	
s descriptive statistics variables u R), SIZE, EXCH, and GDPR in IDR; S	stics presented are average, medi :he whole sample data	
tive statistics. This table reports housand Indonesian Rupiah (IDF	in percentage. Descriptive statis statistics are calculated over t	
Table 2. Descript comma); SP in th	CAPI, and INTR i observation. The	

ouservation. Ine	ב שומוושרורש מוב כמור	niarea over the v	מווטופ אוווטופ ממוח					
						Perce	entile	No. of
Var	Average	Median	Max	Min	Std dev	1%	%66	Obeservations
SP	1.72	0.60	30.95	0.02	3.14	0.03	15.45	3,318
MIM	3.93	3.44	12.50	0.24	2.71	0.24	12.50	3,318
ROE	5.26	4.61	23.28	-33.12	7.79	-33.12	23.28	3,318
NPL	4.35	3.15	37.98	0.32	4.98	0.32	37.98	3,318
LAW	2.33	1.92	11.29	0.02	1.96	0.02	11.29	3,318
DLR	1.32	1.22	2.65	0.91	0.31	0.91	2.65	3,318
LAR	28.38	26.08	61.31	10.54	9.83	10.54	61.31	3,318
SIZE	17.44	17.53	20.94	13.71	1.81	14.06	20.75	3,318
CAPI	14.70	14.13	32.21	-12.27	5.05	5.98	29.25	3,318
EXCH	9.31	9.29	9.63	9.05	0.18	9.05	9.61	3,318
INTR	6.74	6.65	11.75	3.82	1.57	3.83	11.25	3,318
GDPR	14.20	14.51	14.85	13.05	0.65	13.05	14.85	3,318

Table 3. Co	prrelation mo	atrix. This tal	ble reports si	mple bivaria	te correlatio	n statistics (Pearson) be	tween variab	ples used in t	the study. Co	rrelation sta	tistics are
calculated	over the wh	ole sample d	data. Correlat	tions are pre	sented as a	half-triangle	matrix					
	SP	MIN	ROE	NPL	LAW	DLR	LAR	SIZE	CAPI	EXCH	INTR	GDPR
SP	1.00											
MIM	0.12*	1.00										
ROE	0.23*	0.46*	1.00									
NPL	-0.19*	-0.02	-0.32*	1.00								
LAW	0.09*	0.14*	-0.13*	0.45*	1.00							
DLR	-0.06*	0.23*	•60.0	0.25*	0.10*	1.00						
LAR	0.02	0.06*	*60.0	0.07*	-0.02	0.66*	1.00					
SIZE	0.52*	0.12*	0.37*	-0.26*	0.19*	-0.16*	-0.10*	1.00				
CAPI	0.32*	0.13*	-0.03	-0.16^{*}	0.14*	-0.22*	0.03	0.16*	1.00			
EXCH	0.20*	-0.21*	-0.18*	-0.12*	0.14*	-0.40*	-0.19*	0.25*	0.25*	1.00		
INTR	-0.13*	0.19*	0.04*	0.16^{*}	0.19*	0.30*	0.00	-0.18^{*}	-0.21*	-0.27*	1.00	
GDPR	0.19*	-0.26*	-0.10*	-0.19*	-0.15*	-0.43*	0.01	0.26*	0.28*	0.59*	-0.72*	1.00

independence is significantly lower than 0.500. Simple correlation coefficients among SP are high with SIZE, at 0.52, and these statistics provide clues about the possibility of a large and significantly positive correlate with stock price measures. In summary, based on the analysis from the data set for subsequent analysis, the profile of data is not believed to have further negative implications.

Table 4 summarizes the findings of the unit root check of the variables used in the study. Significant rejection of nonstationary null hypothesis is obtained for variables NIM, ROE, NPL, LAW, DLR, LAR, CAPI, and INTR. On the other hand, variables SP, SIZE, EXCH, and GDPR are significantly non-stationary.

Table 5 shows that null hypotheses of no cointegration are not rejected using subgroup and global criteria. Nevertheless we can see later, from ECM model estimates that the ECT is highly significant, indicating that there is an equilibrating mechanism, i.e. cointegration. We suspect that the apparent inalignment is caused by undetected multicollinearity or endogeneity among risk proxies used (credit risk and liquidity).

Finally, as shown in Table 6, the null hypotheses of slope homogeneity are soundly rejected. Therefore, the regressions should be estimated by the mean group type estimator. We do not report the exante cross-section dependence test as the result is similar with post-estimation procedure that strongly concludes the presence of cross-sectional dependence. The presence of cross-sectional dependence then called for the use of cross-section averaging of variables to improve the estimation of equation. With this result, we conclude our estimation technique as a Dynamic Common Correlated Error Estimator with Mean Group variant (DCCE-MG).

3.2. Regression result

The regressions provide sound support for the error-correction model. We find strong and significant evidence of an error-correction mechanism from the regressions. The MG estimate for ECT is -0.3094, which corresponds to a time lag of 3.2 months (=1/0.3094). The CD test yields significant statistics, indicating that cross-sectional dependence occurs and is significant.

In Table 7, ROE, LAW, DLR, and SIZE were positively correlated with SP, but ROE has signicance to SP in short-run equations and long-run equations, while DLR has significance in short-run equations only. Therefore, it seems that the DLR impact on SP is a more short-term reaction rather long-term adjustment.

The findings show that macroeconomic (EXCH, INTR, and GDPR) is negatively correlated with SP and this result is in accordance with the literature from Huy et al. (2020).

3.3. Robustness test

In Table 8, the outcome of ECT estimates might be considered qualitatively equal to regression for the robustness test using SP. The MG and DCCE criteria have an approximate range from -0.3103 to -0.3156, and these ECT estimations correlate with a time lag of 3.17 months (=1/0.3156) to 3.22 months (=1/0.3103). For all models, the CD test yields significant statistics, indicating that cross-sectional dependence occurs and is significant.

According to the results, all independent variables (NIM, ROE, NPL, DLR, and ALR) except LAW in all models were positively correlated with SP and these results are the same as baseline regression. DCCE estimations in models 1 and 2 are preferred and as predicted ROE by the long-run equation in Models 1, 2, and 3 and the short-run equation in Models 1 and 2 was positively correlated and significant to SP. Control variables did not have a significant impact on SP.

In Table 9, the outcome of ECT estimates might be considered qualitatively equal to regression for the robustness test using SP and reveals that SOE banks are the most responsive in the event of

Table 4. Unit root test. This table reports the summary of unit root tests performed on variables used in the study. The report covers (1) name of variables, (2) name of corresponding unit root test method used, and (3) conclusion of the test (reject or cannot reject the null hypothesis). The levels of significance used as a rule for null hypothesis rejection are 1%, 5%, and 10%, respectively

No.	Var	Method	Result
1	SP	PESCADF	Non-stationary at the 1% significance level
2	NIM	PESCADF	Stationary at the 1% significance level
3	ROE	PESCADF	Stationary at the 1% significance level
4	NPL	PESCADF	Stationary at the 1% significance level
5	LAW	PESCADF	Stationary at the 1% significance level
6	DLR	PESCADF	Stationary at the 1% significance level
7	LAR	PESCADF	Stationary at the 1% significance level
8	SIZE	PESCADF	Non-stationary and non- significance level
9	САРІ	PESCADF	Stationary at the 1% significance level
10	ЕХСН	ADF	Non-stationary and non- significance level
11	INTR	ADF	Stationary at the 1% significance level
12	GDPR	ADF	Non-stationary and non- significance level

a stock price shock. Using SP as a stock price gauge, the ECT estimates for SOE banks were -0.4694 or 2.13 months shock adjustment, and these are greater than PRIV banks, which were -0.3015 or 3.32 months shock adjustment. This may be due to SOE banks' extraordinary reputation or credibility, which makes them purchase or release stock easily.

According to the robustness test, the result of long-run and short-run equations, ROE in Model 5a, were positively correlated, significant to SP, and consistent with the baseline regression finding. The short-run equation of CAPI in Model 5b was negatively correlated with and highly significant to SP. This is probably because State-Owned Enterprises (SOEs) have a strong capital and different business model than National Private Bank (PRIV). Customer financing appears to be a focus of SOE banks' activity and the signaling concept appears to work for both SOE and PRIV banks.

ROE and LAW have different algebraic signs, while our initial hypothetical sign of ROE is positive and LAW is negative, so they are still similar to our initial hypothesis. All control variables except CAPI were not significant to SP.

In Table 10, the sample in Model 6b (after April 2009) had lower ECT coefficients than the sample in Model 6a (up to April 2009). Therefore, this estimate is in line with the idea that after the 2008 economic crisis, there was a shock that fell stock prices after April 2009.

ROE by the long-run equation is positive and significant after April 2009, but lost relevance in the sample up to April 2009. The stock price signaling hypothesis appears to be valid just after April 2009. This finding is aligned with Simoens and Vennet, 2020 who concluded that one of

Table 5. Cointegratio and Westerlund's (20 cointegration (using	n test. This table report 113) methods. Cointegra Pt and Pa test statistics	s the results of the coin ttion is based on various s) and (2) subsets cointe	tegration test based or specifications: with SP gration (using Gt and G	 Westerlund (2007), Pe as the dependent varia a) for all specifications 	rrsyn and Westerlund able. Null hypotheses	(2008), and Blomquist are (1) global no
No	Specification	Statistic;	Gt	Ga	Pt	Ρα
1	SP; NIM, ROE, NPL, LAW,	Z-value	1.52	2.50	4.57	3.64
	DLR, LAR, Constant, Lags: 1–4, Leads: 1–4, LR Window: 4, Selection: AIC (Lag: 3.86; Lead: 3.86)	P-value	0.94	66.0	1.00	1.00
2	SP; NIM, ROE, NPL, LAW,	Z-value	-0.53	1.67	1.69	1.87
	DLR LAR, Constant- Trend, Lags: 1-4, Leads: 1-4, LR Window: 4, Selection: AIC (Lag: 4; Lead: 3.86)	P-value	0.30	0.95	96.0	0.97
c	SP; NIM, ROE, NPL, LAW,	Z-value	1.68	1.68	4.23	3.53
	DLR, LAR, Constant, Lags: 1, Leads: 1, LR Window: 1, Bootstrap	Robust P-value	0.84	0.88	0.98	0.99
4	SP; NIM, ROE, NPL, LAW,	Z-value	1.71	2.05	3.70	3.30
	DLR, LAR, Constant, Lags: 2, Leads: 2, LR Window: 2, Bootstrap	Robust P-value	0.70	0.89	0.93	0.95
5	SP; NIM, ROE, NPL, LAW,	Z-value	2.15	2.90	5.06	4.05
	DLR, LAR, Constant, Lags: 3, Leads: 3, LR Window: 3, Bootstrap	Robust P-value	0.80	0.97	0.97	0.95
9	SP; NIM, ROE, NPL, LAW,	Z-value	3.18	4.50	4.70	5.02
	DLR, LAR, Constant, Lags: 4, Leads: 4, LR Window: 4, Bootstrap	Robust P-value	0.92	1.00	0.98	0.99
7	SP; NIM, ROE, NPL, LAW,	Z-value	0.51	0.98	1.52	1.43
	DLR, LAR, Constant, Trend, Lags: 1, Leads: 1, LR Window: 1, Bootstrap	Robust P-value	0.41	0.57	0.77	0.77
						(Continued)

I nnie J. (continueu)						
No	Specification	Statistic;	Gt	Ga	Pt	Pa
8 SP; N	JIM, ROE, NPL, LAW,	Z-value	0.37	1.41	1.63	2.06
DLR, Trenc LR W	LAR, Constant, d, Lags: 2, Leads: 2, /indow: 2, Bootstrap	Robust P-value	0.29	0.66	0.71	0.81
9 SP; N	JIM, ROE, NPL, LAW,	Z-value	0.65	2.28	3.24	3.23
DLR, Tren. LR W	LAR, Constant, d, Lags: 3, Leads: 3, /indow: 3, Bootstrap	Robust P-value	0.32	0.72	0.77	0.75
10 SP; N	JIM, ROE, NPL, LAW,	Z-value	1.19	4.61	1.87	4.92
DLR I Trenc LR W	LAR, Constant, d, Lags: 4, Leads: 4, /indow: 4, Bootstrap	Robust P-value	0.24	1.00	0.53	0.99

Table 6. Slope Heterogeneity Test. This table reports the slope heterogeneity test using the method proposed by Pesaran and Yamagata (2008) and Blomquist and Westerlund (2013). There are several specifications based on (1) whether SP is a dependent variable, (2) inclusion of cross-section average, and (3) heteroscedasticity treatment. Null hypotheses are slope homogeneity for all specifications

		Unadjust	ed Delta	Adjuste	d Delta
No	Specification	Stats.	p-value	Stats.	p-value
1	SP; NIM, ROE, NPL LAW DLR, LAR SIZE CAPI, EXCH, INTR, GDPR; HAC	65.28	0.00	68.14	0.00
2	SP; NIM, ROE, NPL, LAW, DLR, LAR,SIZE, CAPI, EXCH. INTR, GDPR; Cross Section Average, CR Lags = 0, HAC	13.70	0.00	14.35	0.00
3	SP; NIM. ROE. NPL. LAW. DLR. LAR. SIZE. CAPI EXCH INTR GDPR; Cross- Section Average, CR Lags = 1, HAC	45.10	0.00	47.48	0.00
4	SP; NIM, ROE, NPL, LAW, DLR, LAR, SIZE, CAPI, EXCH, INTR, GDPR; Cross Section Average, CR Lags = 2, HAC	81.84	0.00	86.67	0.00
5	SP; NIM, ROE NPL, LAW, DLR, LAR, SIZE, CAPI, EXCH, INTR, GDPR; Cross- Section Average, CR Lags = 3, HAC	22.10	0.00	23.58	0.00
6	SP; NIM, ROE, NPL LAW DLR, LAR SIZE CAPI, EXCH, INTR, GDPR; Cross Section Average, CR Lags = 4, HAC	17.40	0.00	18.74	0.00

the main drivers of bank M/B ratio from 2007 to 2017 was profitability, while cost efficiency improvement was a key driver of better profitability following the crisis.

LAW negatively correlated with SP is equal to the initial hypothetical sign, and according to the literature from Phung et al. (2021), credit risk negatively correlate with bank efficiency and ultimately negatively correlate with stock prices, which is similar towith Abbas et al. (2019), The impact of liquidity on profitability positively correlates, while the credit risk has a negative and statistically significant effect on bank profitability. All control variables were not significant.

Table 7. Baseliene regression result. This table reports estimations of the baseline model with a dependent variable. The table presents the estimated coefficients and standard error. The regression model corresponds to specific specifications inclusion of constant and application of cross-section average (DCCE class estimators). Statistical significance is denoted as * at the 10% level, ** at the 5% level, and *** at the 1% level

Variables		Baseline Ro	egression
Dep. Var: D.SP		Short Run	Long Run
ROE	Coefficient	0.0047**	0.0152**
	Standard error	0.0023	0.0075
LAW	Coefficient	0.0083	0.0432
	Standard error	0.0161	0.0514
DLR	Coefficient	0.1960*	0.1768
	Standard error	0.1091	0.2674
SIZE	cCoefficient	0.1139	
	Standard error	0.1527	
CAPI	Coefficient	-0.0035	
	Standard error	0.0053	
EXCH	Coefficient	-0.1553	
	Standard error	0.1613	
INTR	Coefficient	-0.0021	
	Standard error	0.0085	
GDPR	Coefficient	-0.0308	
	Standard error	0.0293	
ECT	Coefficient		-0.3094***
	Standard error		0.0302
Constant	Coefficient		-1.4376
	Standard error		6.6149
ECM Specifications			
Cross-Section Averages an	d Constant		
Observations		3,297	
F Stat		3.42***	
CD Test		9.71***	

In general, NPL and LAW negatively correlated with SP and are equal to the initial hypothetical sign. According to the literature from Phung et al. (2021), non-performing loans negatively correlate with bank efficiency and ultimately negatively correlate with stock prices

4. Conclusion

Our paper has shown that profitability is the main driver of bank stock prices in Indonesia but have to be viewed with their risk features. Nevertheless, the conventional risk measures do not seem to have a significant effect. This result has important consequences for bank management, regulators, and investors, in which perhaps new more meaningful risk measures should be developed, tested, and imposed.

We have modelled the empirical relationship of bank stock price with its selected fundamental (profitability, credit risk, and liquidity risk) using DCCE in ECM format. The empirical design is then applied to a monthly panel data set composed of 21 banks from period July 2006 to August 2019 (3.318 bank month observations). Our empirical design is quite well supported. We found a highly significant coefficient of ECT between 0.31 and 0.38, corresponding to 2.62 to 3.22 months of equilibrating time. ECT estimates are proven to be robust under various checking schemes.

Table 8. Robu estimated coe specifications Statistical sig	stness test with efficients and sta inclusion of con nificance is denc	other variables. andard error. Eac istant and applice bted as * at the 1	This table report h regression mo ation of cross-se 10% level, ** at t	ts estimations of del (denoted by ction average, b he 5% level. and	f the baseline m a number in the y alternating pr H *** at the 1%	iodel with the de e first line of ead roxies in the vari level	pendent variabl :h column headi ables of interes'	e. The table pres ng) corresponds t (DCCE class est	sents the to specific imators).
Variables		Ром	lel 1	poM	el 2	Mod	el 3	Mod	el 4
Dep. Var: D.SP		Short Run	Long-Run	Short-Run	Long-Run	Short-Run	Long-Run	Short-Run	Long-Run
MIM	Coeff.							0.0140	0.0051
	SE							0.0142	0.0430
ROE	Coeff.	0.0047**	0.0146*	0.0038*	0.0130**	0.0036	0.0133**		
	SE	0.0023	0.0079	0.0019	0.0059	0.0022	0.0063		
NPL	Coeff.	0.0110	0.0712	0.0109	0.0574			0.0100	0.0635
	SE	0.0282	0.1095	0.0213	0.0981			0.0281	0.1139
LAW	Coeff.					-0.0002	0.0217		
	SE					0.0103	0.0580		
DLR	Coeff.	0.1471	0.0886					0.0808	0.0014
	SE	0.0931	0.2307					0.0785	0.1924
LAR	Coeff.			0.0037	0.0035	0.0043	0.0081		
	SE			0.0026	0.0056	0.0028	0.0070		
SIZE	Coeff.	0.0332		-0.0103		-0.0169		0.0605	
	SE	0.1928		0.2153		0.1470		0.1932	
CAPI	Coeff.	-0.0062		-0.0041		-0.0012		-0.0078	
	SE	0.0081		0.0083		0.0052		0.0083	
EXCH	Coeff.	-0.1200		-0.1482		-0.1686		-0.1080	
									(Continued)

Table 8. (Conti	nued)								
Variables		Mod€	1 1 Ie	₩ode	il 2	Mode	el 3	Mode	el 4
Dep. Var: D.SP		Short Run	Long-Run	Short-Run	Long-Run	Short-Run	Long-Run	Short-Run	Long-Run
	SE	0.1476		0.1393		0.1589		0.1594	
INTR	Coeff.	-0.0003		-0.0024		-0.0049		-0.0014	
	SE	0.0116		0.0118		0.0087		0.0115	
GDPR	Coeff.	-0.0112		-0.0349		-0.0144		-0.0060	
	SE	0.0182		0.0384		0.0273		0.0230	
ECT	Coeff.		-0.3156***		-0.3151***		-0.3103***		-0.3151***
	SE		0.0311		0.0059		0.0303		0.0321
Constant	Coeff.		-0.4497		2.8842		4.8027		-2.5225
	SE		7.4836		9.0746		7.5699		7.8377
ECM Specifications									
Observations		3,25	۲۴	3,25		3,29	26	3,29	26
F Stat		3.86	***	3.83	***	3.47	***	3.84	***
CD Test		8.37	***	8.56	***	8.90	***	8.45	***

Table 9. Robustness test—Bank type. This table reports estimation of the robustness model: bank type regressions with the dependent variable using the DCCE technique with ECM format. The table presents the estimated coefficients and standard error. Each regression model (denoted by a number in the second line of each column heading) corresponds to specific specifications described by inclusion of constant and cross-section average. Estimations are applied to subsamples based on bank types ("SOE" and "PRIV"). Statistical significance is denoted as * at the 10% level, ** at the 5% level, and *** at the 1% level

Variables		Model 5a		Model 5b	
Dep. Var: D.SP		National Private Bank (PRIV)		State-Owned Enterprises (SOE)	
		Short-Run	Long-Run	Short-Run	Long-Run
L.ROE	Coefficient	0.0074**	0.0228**	-0.0013	-0.0013
	Standard error	0.0036	0.0107	0.0022	0.0039
L.LAW	Coefficient	0.0116	0.0470	-0.0081	0.0237
	Standard error	0.0227	0.0495	0.0128	0.0232
L.DLR	Coefficient	0.0993	0.1598	0.2278	0.0225
	Standard error	0.0689	0.2644	0.2111	0.0957
SIZE	Ccoefficient	0.1990		-0.1870	
	Standard error	0.2010		0.4370	
CAPI	Coefficient	-0.0023		-0.0033*	**
	Standard error	0.0034		0.0011	
EXCH	Coefficient	-0.2072		0.1345	
	Standard error	0.1656		0.8554	
INTR	Coefficient	-0.0017		-0.0007	
	Standard error	0.0120		0.0159	
GDPR	Coefficient	-0.0531		0.0589	
	Standard error	0.0467		0.0895	
ECT	Coefficient		-0.3015***	k	-0.4694***
	Standard error		0.0299		0.1568
Constant	Coefficient		-3.5837		0.1342
	Standard error		9.8253		2.7221
ECM Specifications					
Cross-Section Averages and Constant					
Observations		2,512		785	
F Stat		3.34***		22.97***	
CD Test		5.41***		-2.20**	

ROE has positive and statistically significant correlation with SP, and DLR has a positive and quite significant correlation with SP. Aligned with our initial hypothesis, profitability (NIM and ROE) and liquidity risk (DLR and ALR) consistent positively correlated with SP and these findings are consistent with the study by Pennacchi and Santos (2021), "when the bank wants to create a rosy picture of its performance, ROE i selected." This option is backed by the widespread use of ROE in bank compensation contracts and investor communications and consistent too with the study by Abbas et al. (2019); the impact of liquidity on profitability positively correlates, while the credit risk negatively and significantly impacts the profitability. Results from the regression for credit risk, LAW, have different algrebraic signs but NPL consistently positively correlated with SP, and this finding is consistent with the study by Madugu et al. (2020), credit risk positive (NPL) and

Table 10. Robustness test—Period. This table reports estimation of robustness check regression with the dependent variable using the DCCE technique with ECM format. The table presents the estimated coefficients and standard error. Each regression model (denoted by a number in the first line of each column heading) corresponds to specific specifications described by inclusion of constant and cross-section average. Estimations are applied to split samples ("Up to April 2009" and "after April 2009"). Statistical significance is denoted by * at the 10% level, ** at the 5% level, and *** at the 1% level

Variables		Model 6a		Model 6b	
Dep. Var: D.SP		Up to April 2009		After April 2009	
		Short-Run	Long-Run	Short-Run	Long-Run
L.ROE	Coefficient	0.0043	0.0048	0.0031	0.0125**
	Standard error	0.0062	0.0082	0.0022	0.0057
L.LAW	Coefficient	-0.0077	0.0304	-0.0020	-0.0244
	Standard error	0.0354	0.0983	0.0125	0.0451
L.DLR	Coefficient	0.1359	-0.1322	0.1325	-0.0894
	Standard error	0.1121	0.1437	0.1281	0.5317
SIZE	Coefficient	-0.0812		0.2208	
	Standard error	0.3093		0.3047	
CAPI	Coefficient	-0.0079		-0.0005	
	Standard error	0.0155		0.0069	
EXCH	Coefficient	-0.0507		-0.1005	
	Standard error	0.2657		0.2033	
INTR	Coefficient	-0.0119		-0.0100	
	Standard error	0.0123		0.0138	
GDPR	Coefficient	-0.1073		-0.0177	
	Standard error	0.3274		0.0400	
ECT	Coefficient		-0.8606***	ŧ	-0.3804***
	Standard error		0.0750		0.0282
Constant	Coefficient		12.2981		-7.0009
	Standard error		9.5759		11.4948
ECM Specifications					
Cross-Section Averages and Constant					
Observations		693		2,604	
F Stat		9.28***		3.08***	
CD Test		-0.02		7.00***	

significant effect on profitability. Robustness check performed by splitting samples: before crisis and after crisis provide support for the mentioned results.

Nevertheless, evidence of no cointegration found by an earlier stand-alone test should give us a note. There may be a undetected multicollinear structure in risk proxies that weakens the inference. This alignment could call for a revisit to panel cointegration test methodology since by theory, the results from ECT estimates and cointegration test should confirm each other. Therefore, future research should address this issue, finding credit risk and liquidity proxies that are orthogonal to each other while simultaneously being a correct measure. Our results suggest that profitability is the main driver of stock prices in Indonesia but have to be viewed with risk implications. Nevertheless, the conventional risk measures do not seem to have a significant effect. This result has important consequences for bank management, regulators, and investors; perhaps new more impactfull risk measure should be developed, tested, and imposed.

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Notes

 We also think Interest risk, arising from duration mismatch, as an important key risk, however, since we do not have access to the required data, hence, we refrain to include this risk in our study but acknowledge it as our study limitation.

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