

Gauriot, Romain; Page, Lionel; Wooders, John

Article

Expertise, gender, and equilibrium play

Quantitative Economics

Provided in Cooperation with:

The Econometric Society

Suggested Citation: Gauriot, Romain; Page, Lionel; Wooders, John (2023) : Expertise, gender, and equilibrium play, Quantitative Economics, ISSN 1759-7331, The Econometric Society, New Haven, CT, Vol. 14, Iss. 3, pp. 981-1020, <https://doi.org/10.3982/QE1563>

This Version is available at:

<https://hdl.handle.net/10419/296342>

Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

Terms of use:

Documents in EconStor may be saved and copied for your personal and scholarly purposes.

You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.

If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.



<https://creativecommons.org/licenses/by-nc/4.0/>

Supplement to “Expertise, gender, and equilibrium play”

(*Quantitative Economics*, Vol. 14, No. 3, July 2023, 981–1020)

ROMAIN GAURIOT

Department of Economics, Deakin University

LIONEL PAGE

School of Economics, University of Queensland

JOHN WOODERS

Division of Social Science, New York University Abu Dhabi and the Center for Behavioral Institutional Design

APPENDIX D: ROBUSTNESS OF POWER SIMULATIONS

As a robustness check, here we reproduce the simulation results reported in Section 6, but where now the simulated data matches the characteristics of the observed data, point game by point game, rather than just in aggregate. Specifically, if point game i has n_R^i serves to the right, n_L^i serves to the left, and an empirical winning frequency of \hat{p}^i , then the simulated data for point game i has n_R^i serves to the right, n_L^i serves to the left, and the probability of winning a point is \hat{p}^i for serves in each direction (and hence the null hypothesis that $p_L^i = p_R^i$ is true). The number of winning serves to the right and left are therefore distributed, respectively, $B(n_R^i, \hat{p}^i)$ and $B(n_L^i, \hat{p}^i)$ in the simulated data for point game i .

The power of our test

The subsection “The power of our test” in Section 6 provided the power functions for the Pearson joint test and the KS tests based on the Pearson p -values and the Fisher exact t -values. It demonstrated that for “small” samples of 40 point games, the test based on the t -values was substantially more powerful than the other two. In addition, for “large” samples of 7000 point games, the test based on the t -values was especially powerful—the joint null hypothesis of equality of winning probabilities is almost surely rejected for even small departures from equilibrium play.

Table D.1 is the analogue Table 6. Comparing to the two tables reveals that the KS test based on the t ’s has similar power when the simulated data matches the characteristics

Romain Gauriot: romain.gauriot@deakin.edu.au

Lionel Page: lionel.page@uq.edu.au

John Wooders: john.wooders@nyu.edu

TABLE D.1. Rejection rate for H_0 at the 5% level, $N = 7198$.

True θ	KS based on t 's	KS based on p 's	Pearson joint test
0.65	0.834	1	0.746
0.66	0.212	1	0.716
2/3	0.051	1	0.726
0.67	0.093	1	0.726
0.68	0.675	1	0.764

of the Hawk-Eye data. Table D1 shows that the (true) joint null hypothesis of equality of winning probabilities is rejected, at the 5% significant level, for sure by the KS test based on the p 's and it is rejected with probability 0.726 by the Pearson joint test. These results reaffirm our conclusion that these tests are not valid for large samples.

Figure D.1 is the analogue to Figure 14 and shows that the power functions in Figure 14 are largely unchanged when the data is simulated (under the null hypothesis) to match the characteristic of the WW data (Figure D.1(a)) or the Hawk-Eye data (Figure D.1(b)).

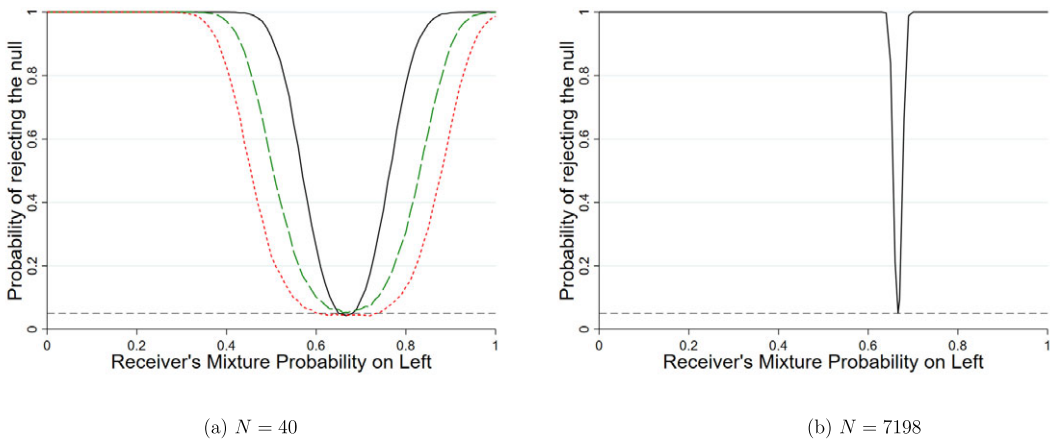


FIGURE D.1. Power functions for KS test based on t -values (black), p -values (red), and Pearson joint (green).

APPENDIX E: BALL BOUNCES

Figure E.1 below shows actual and imputed ball bounces for male second serves from the deuce court.

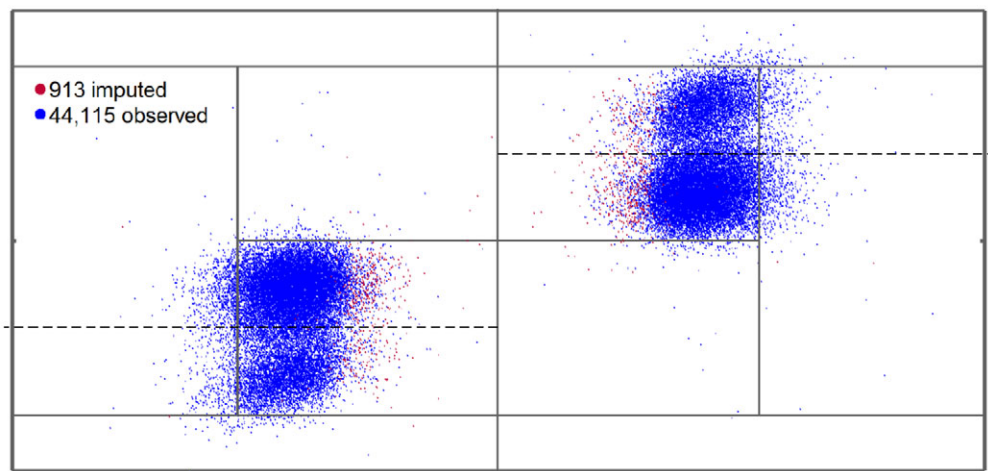


FIGURE E.1. Ball bounces for deuce-court second serves by men.

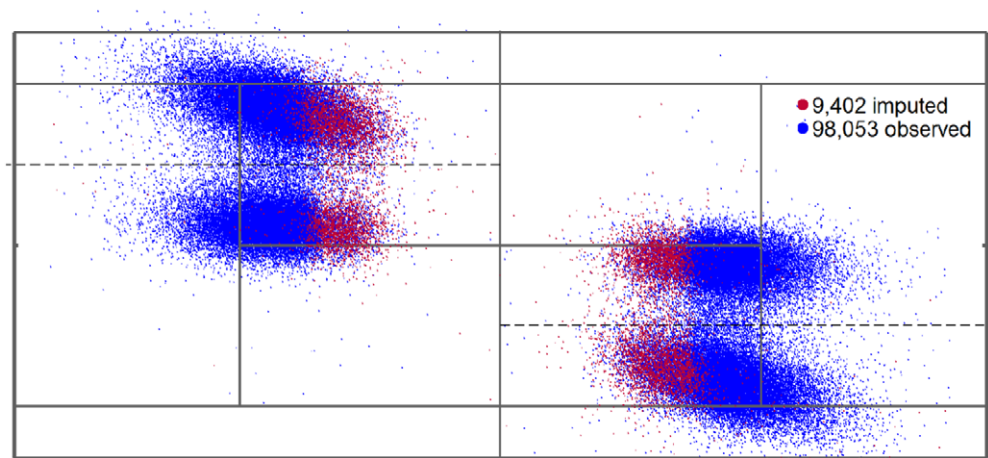


FIGURE E.2. Ball bounces for ad-court first serves by men.

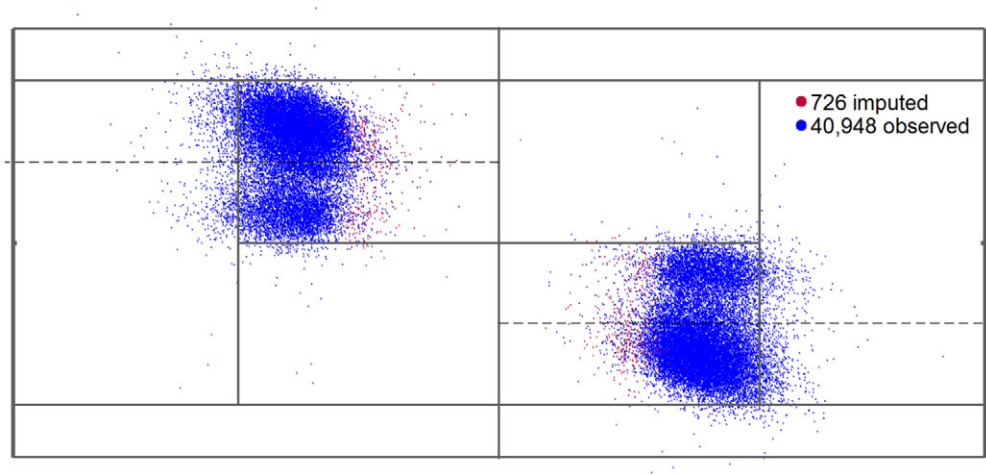


FIGURE E.3. Ball bounces for ad-court second serves by men.

Ball bounces for first and second serves by women are below.

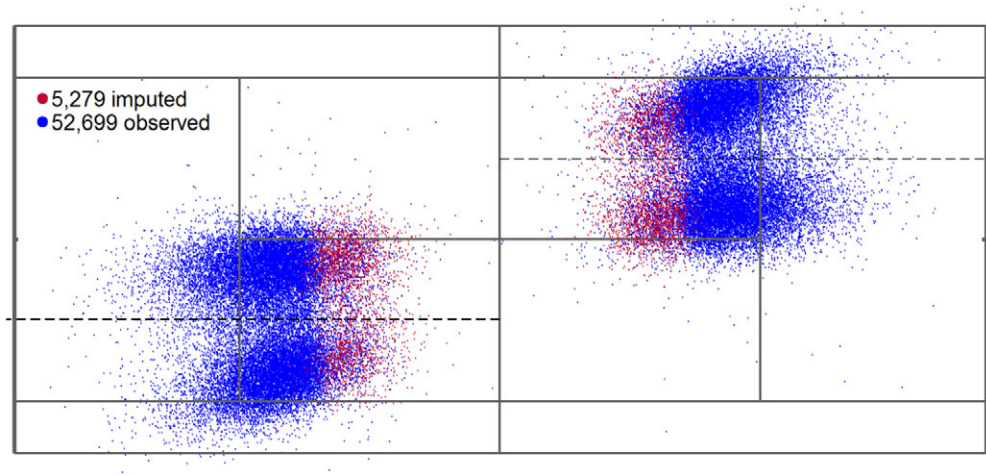


FIGURE E.4. Ball bounces for deuce-court first serves by women.

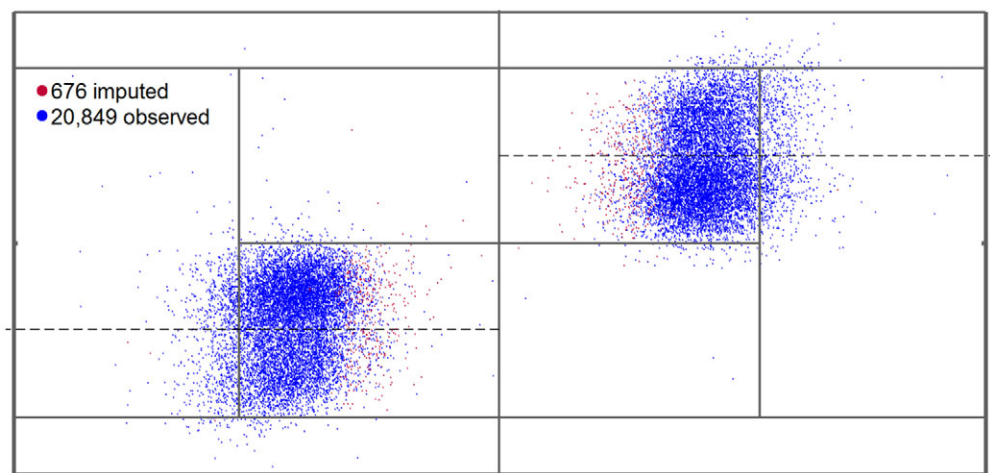


FIGURE E.5. Ball bounces for deuce-court second serves by women.

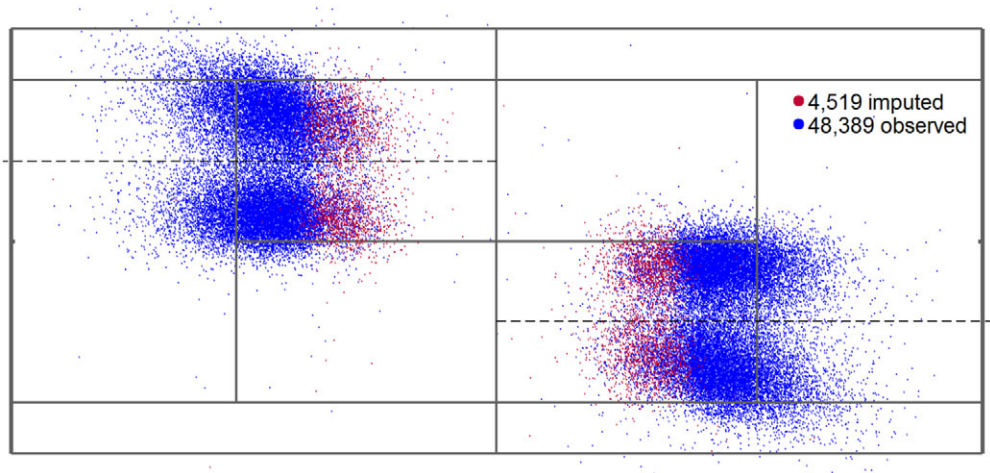


FIGURE E.6. Ball bounces for ad-court first serves by women.

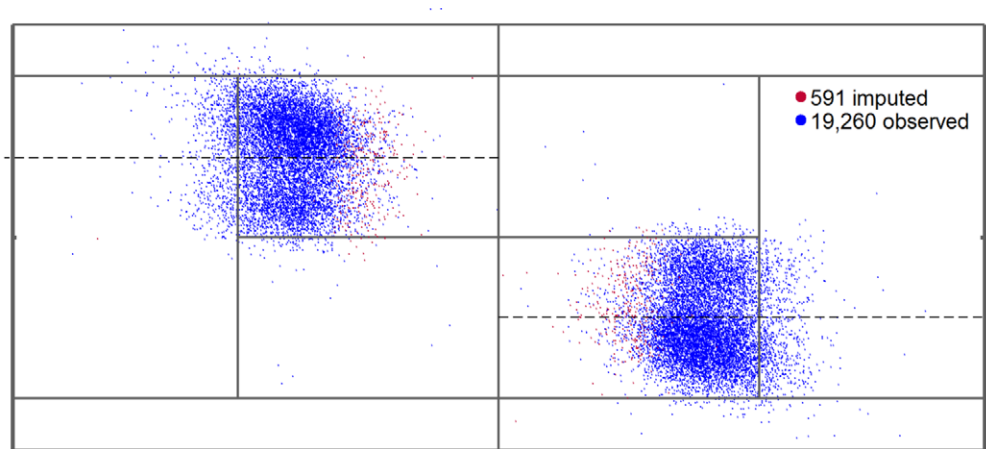


FIGURE E.7. Ball bounces for ad-court second serves by women.

Co-editor Peter Arcidiacono handled this manuscript.

Manuscript received 2 March, 2020; final version accepted 30 January, 2023; available online 14 March, 2023.