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Gächter, Simon; Lee, Kyeongtae; Sefton, Martin; Weber, Till O.

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The Role of Payoff Parameters for Cooperation in the One-Shot Prisoner's Dilemma

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Simon Gächter

University of Nottingham, IZA and CESifo

Kyeongtae Lee

Bank of Korea

Martin Sefton

University of Nottingham

Till O. Weber

Newcastle University

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ABSTRACT

The Role of Payoff Parameters for Cooperation in the One-Shot Prisoner's Dilemma

The prisoner's dilemma (PD) is arguably the most important model of social dilemmas, but our knowledge about how a PD's material payoff structure affects cooperation is incomplete. In this paper we investigate the effect of variation in material payoffs on cooperation, focusing on one-shot PD games where efficiency requires mutual cooperation. We report results from three experiments (N = 1,993): in a preliminary experiment, we vary the payoffs over a large range. In our first main experiment (Study 1), we present a novel design that varies payoffs orthogonally in a within-subjects design. Our second main experiment, Study 2, investigates the orthogonal variation of payoffs in a between-subjects design. In a complementary analysis we also study the closely related payoff indices of normalized loss and gain, and the K-index. The most robust finding of our experiments and the complementary analyses is that cooperation in a PD increases with the gains of mutual cooperation over mutual defection.

JEL Classification: A13, C9⁻

Keywords: prisoner's dilemma, cooperation, payoff parameters,

temptation, risk, efficiency, normalized gain, normalized loss,

K-index, experiments

Corresponding author:

Simon Gächter Centre for Decision Research and Experimental Economics (CeDEx) University of Nottingham Nottingham NG7 2QX United Kingdom

E-mail: simon.gaechter@nottingham.ac.uk

1. Introduction

In many economic and social environments there is a conflict between individual and collective interests. The simplest model to represent such a conflict is the Prisoner's Dilemma (PD) and so it plays an important role in the behavioral sciences, where the PD is the topic of a vast literature in economics, sociology, political science, and social psychology. There is extensive evidence of cooperation in experimental PDs, and cooperation is observed even in carefully controlled anonymous one-shot interactions where participants have a real material incentive to defect (e.g., Cooper et al. (1996); Frank et al. (1993); Mengel (2018); Embrey et al. (2018); Dal Bó and Fréchette (2018)). The literature has studied a wide variety of factors that affect cooperation (see, e.g., Balliet et al. (2009); Balliet (2010); Van Lange et al. (2014)), but perhaps from an economics perspective the most fundamental factor to consider is the material payoff structure. If players would be solely motivated by material payoffs, defecting would be a dominant strategy and the structure of payoffs – the relative size of payoffs in the PD – would not matter.

The fact that people sometimes cooperate in anonymous one-shot PDs violates the assumption that people always maximize material payoffs. Given this observation, we ask the most basic question, which we will make precise below: which features of the payoff structure explain cooperation? As we discuss in detail in Section 2, a surprisingly small literature has studied this question and a robust result has yet to emerge. Our contribution is to provide, across three experiments, a systematic analysis of the role of the material payoff structure for cooperation in one-shot PDs.

Our experiments are based on games in which two participants simultaneously choose to either 'Cooperate' or 'Defect' and their choices translate into money earnings as shown in Table 1.

Table 1. The Prisoner's Dilemma game.

| | Cooperate | Defect |
|-----------|-----------|--------|
| Cooperate | R, R | S, T |
| Defect | T, S | P, P |

Notes: T > R > P > S. Row's payoff is given by the first entry in each cell.

¹ Cooperation is also observed in repeated PD games that allow for strategic motives to cooperate (see, e.g., Embrey et al. (2018)). For a discussion of cooperation in finitely and infinitely repeated PD game experiments, see Mengel (2018) and Dal Bó and Fréchette (2018), respectively.

We refer to the entries in Table 1 as payoffs, but to be clear they are the material payoffs resulting from their decisions and we make no claim about how they are related to utility more broadly construed. Following Rapoport and Chammah (1965) we choose the payoffs to satisfy the PD condition T > R > P > S. Thus, participants earn more from mutual cooperation than from mutual defection (R > P). However, cooperation is a 'risky' choice that makes the participant vulnerable to being exploited by a defector (P > S). Additionally, each participant is 'tempted' to choose defection as it increases her earnings against a cooperator (T > R). The PD condition ensures that the dominant strategy for money-maximizing participants is to defect. Rapoport and Chammah (1965) impose a second condition, 2R > T + S, to ensure that mutual cooperation maximizes combined earnings. The remainder of the paper focuses on one-shot PDs that satisfy both conditions.

Our goal in this paper is to investigate the role of the relative size of payoffs in Table 1 for cooperation. We are interested in three natural—*ceteris paribus*—payoff comparisons that capture the three sources of incentives alluded to in the previous paragraph:

First, a row player who assumes column player plays Cooperate gets a payoff increase of T - R from defecting rather than cooperating. While a selfish player would defect since T - R is positive, more generally a player who cares about own payoffs but trades this off against other considerations would have an increased incentive to defect as T - R increases.² Thus, the higher T is relative to R, the higher the *temptation* to defect, holding all other payoffs constant.

Second, a row player who assumes column player plays Defect gets a payoff increase of P-S from defecting rather than cooperating. Equivalently, a player who cooperates *risks* getting S rather than the payoff of P they would have got from defecting. Again, there is an increased incentive to defect and a decreased incentive to cooperate as P-S increases. These two payoff comparisons are based on a player's interest in own payoff.

Third, it is also possible that players are motivated by collective interests, and so we consider a further payoff comparison, whereby players might also be more likely to cooperate the greater the payoff from mutual cooperation, R, is relative to the payoff from mutual defection, P, that is, the greater the *efficiency* of cooperation R - P.

We express these *ceteris paribus* payoff comparisons as percentage changes, using Mengel's (2018) payoff indices TEMPT $\equiv \frac{T-R}{T}$; RISK $\equiv \frac{P-S}{P}$; and EFF $\equiv \frac{R-P}{R}$. A property of the RISK, TEMPT, EFF indices is that they are invariant to multiplying the game's payoff

² These other considerations could, for example, reflect other-regarding concerns, such as utility derived from the payoffs of others.

matrix by a factor, for example, when using varying exchange rates across different subject pools. However, they are not invariant to adding a constant (e.g., in case of differing show-up fees). While this does not concern the within-subject pool investigation of the relative explanatory power of the three indices, a careful comparison across studies with varying show-up fees might warrant the use of normalized indices (see Section 5).

In the previous literature, several payoff indices have been proposed to predict the degree of cooperation in PDs (see Murnighan and Roth (1983)). Perhaps best known is Rapoport (1967)'s K-index ($\frac{R-P}{T-S}$) which is defined as the gains from mutual cooperation over mutual defection, (R-P), normalized by the payoff range (T-S). The K-index condenses a game's incentives into a single index based on all four elements of the payoff matrix. This can be viewed as a parsimonious prediction of how likely cooperation will be for a given payoff structure, but it has the disadvantage that PD games with very different incentives in terms of RISK, TEMPT and EFF may have the same K-index. In fact, several studies report varying rates of cooperation across PD games with different payoffs but the same K-index (e.g., Moisan et al. (2018)). Our approach, based on Mengel's indices, is not to predict the overall rate of cooperation in a game (that will no doubt depend on all four material payoffs, plus a host of other factors) but rather to examine the *ceteris paribus* effects of changes in particular incentives, allowing a more nuanced examination of the effects of payoff variation.

Our experiments are motivated by several observations about the previous literature and a preliminary experiment (which we will discuss in Section 2). First, the earliest studies and most of the subsequent research has examined payoff effects in the context of *repeated* PDs. Here, of course, players may have strategic reasons to cooperate, at least in early periods. This in turn complicates the interpretation of payoff indices as measuring incentives to defect. For example, for a given payoff matrix the incentive to defect differs according to whether a player is making a choice in the first or the last period.

Second, there are surprisingly few studies that have examined the effect of controlled payoff variation on cooperation in *one-shot* PDs and these offer an incomplete account of the role of material incentives for several reasons. Most of these studies vary more than one payoff index simultaneously across treatments and therefore cannot provide clear evidence on the relative effect size across the payoff indices.

Furthermore, most of these studies eliminate strategic reasons to cooperate by randomly matching participants across periods, but by allowing feedback between games they do allow for learning effects. For example, even if a participant plays against different participants across

periods, the experience of being defected on in early periods may shape a participant's willingness to cooperate in later periods.

Most relevant to our research is Mengel (2018). While few experiments examine controlled variation in payoffs, payoffs do differ considerably across studies and she takes advantage of this variation to conduct a meta-analysis of the roles of RISK and TEMPT, controlling for EFF. For one-shot games Mengel finds that RISK best explains variation in cooperation rates and TEMPT has no explanatory power after controlling for RISK and EFF. However, her meta-analysis includes games that do not meet the Rapoport and Chammah (1965) PD conditions of T > R > P > S and 2R > T + S. As we show in Section 2.2, Mengel's result does not hold when imposing the PD conditions. In the restricted sample satisfying both conditions neither RISK nor TEMPT has a significant effect on cooperation after controlling for efficiency. Moreover, Mengel's study is based on data from experiments that vary in many potentially important procedural variables, as well as in the payoffs they use, and so identifying the effect of payoff variation requires that these other procedural variables do not vary systematically with payoffs, or that they are adequately controlled for. In our experiments we vary payoffs systematically across treatments within a fixed design, offering an opportunity to corroborate (or not) Mengel's results via controlled experimental analysis.

We conduct a preliminary experiment and two new studies motivated by Mengel's results and those of our preliminary experiment. For our preliminary experiment, we run a lab experiment in which participants played 15 one-shot games with varying payoffs in a within-subject design. Payoffs were chosen to meet our two PD conditions while aiming for large variation in the RISK, TEMPT and EFF indices, resembling the variation across the studies that entered Mengel's meta-analysis. Despite wide variation in these payoff indices, we find no evidence that cooperation is systematically related to RISK. We find that cooperation is significantly higher when EFF is higher, and we also find some suggestive evidence that cooperation decreases with TEMPT. However, this design includes only a few instances where one index varies while the other two indices are held constant.

In our first main experiment, called Study 1, we vary RISK, TEMPT and EFF *orthogonally* across eight games that meet the two PD conditions. This allows us to conduct a clean test of the effect of changing one index while holding constant the remaining two. Again, we employ a within-subject design in which participants make decisions in all eight games. We recruit participants from two different subject pools. Our first subject pool is comprised of university student participants, as in most of the studies that motivated our experiment. Our second subject

pool consists of workers on the Amazon Mechanical Turk (AMT) platform, which constitutes a more diverse subject pool regarding age, income, and education (e.g., Arechar et al. (2018); Snowberg and Yariv (2021)). Previous studies have found that cooperation varies systematically with demographic characteristics. For instance, older people tend to cooperate more than the young (e.g., Gächter and Herrmann (2011); List (2004); Matsumoto et al. (2016); Praxmarer et al. (2024)). Comparing subject pools allows us to test whether results based on student samples are transferable to a more diverse and, on average, older and presumably more cooperative population. In neither subject pool do we find any evidence that cooperation varies systematically with RISK. In contrast, cooperation decreases significantly with TEMPT and increases significantly with EFF in both subject pools.

A potential criticism of Study 1 is that the within-subject design allows for learning through enhanced experience in game play or induces an experimenter demand effect whereby participants might feel compelled to condition their action on the payoffs as these are the only things changing across the rounds. In our second main experiment, Study 2, we address this criticism by conducting a between-subject experiment using the same games as in Study 1 and, as far as possible, the same instructions and procedures. We recruit participants from the AMT platform. Participants play a single one-shot PD game, where the game is randomly drawn from one of the eight games used in Study 1. We find that cooperation is significantly higher when EFF is higher, whereas we do not find significant effects of RISK and TEMPT on cooperation.

Taken together, our experiments suggest that, in one-shot PDs where mutual cooperation maximizes social welfare, increasing EFF has a robust and positive impact on cooperation whereas decreasing RISK does not significantly enhance cooperation. Increasing TEMPT has the most detrimental effect on cooperation, but only when participants experience multiple games where payoffs vary. Complementary analyses with the frequently used indices normalized loss, normalized gain, and the K-index, which are related to our indices RISK, TEMPT, and EFF, respectively, support our main conclusion: across all our experiments and subject pools, cooperation in the prisoner's dilemma increases with EFF.

The remainder of this study is organized as follows. Section 2 reviews the related literature and preliminary experiment. Section 3 introduces the design, and Section 4 discusses the results, of our main experimental studies. Section 5 provides evidence on the related indices. In Section 6 we offer a short and tentative discussion of potential explanations of our results. Section 7 concludes.

2. Related literature and some preliminary evidence

There is a vast experimental literature on PDs (for surveys see Balliet et al. (2009); Van Lange et al. (2014)). However, the very first published paper on PD experiments (Flood (1958)), the early work of Rapoport and Chammah (1965), and much of the subsequent experimental literature, has studied repeated PDs. The repeated PD offers a rich environment to study strategic behavior, but a complicated one in which to study the role of payoff structure for cooperation. Embrey et al. (2018) and Mengel (2018) discuss the effect of payoffs on cooperation in finitely repeated PDs. The role of incentives, unconfounded with strategic incentives, is laid bare in the one-shot PD. In the one-shot PD players have a dominant strategy to defect, but nevertheless cooperation is often observed. Many studies have investigated factors promoting cooperation (see, for example, Sally (1995) and Balliet (2010), which survey the role of communication) but there are surprisingly few studies that implement *controlled payoff variation* in the basic one-shot PD. We discuss these in Section 2.1. Of course, payoffs vary greatly across studies, and so Mengel (2018) uses a meta-analysis to study the effect of payoff indices on cooperation. We discuss Mengel's study in Section 2.2.

2.1. Experiments varying payoff parameters

To our knowledge, seven experimental studies examined the effect of controlled payoff variation on cooperation in prisoner's dilemmas. Charness et al. (2016) conducted a one-shot PD between-subject experiment varying R across four treatments. They found that average cooperation rates increase with R. However, note that both EFF and TEMPT change as R changes. Therefore, we cannot say whether increasing R increases cooperation because it increases efficiency, or decreases temptation, or both. Our experiments will allow us to separately identify the effects of EFF and TEMPT on cooperation.

Six studies implemented within-subject experiments where participants played multiple prisoner's dilemma games with varying payoffs. Engel and Zhurakhovska (2016) studied 11 one-shot PDs where *P* varied across games and *T*, *S* and *R* were held constant. Each participant played all 11 PDs with no feedback between games. The authors found that cooperation decreases as *P* increases. Note, however, that this varies RISK and EFF simultaneously across games, and the observed decrease in cooperation may be due to either increasing RISK, or decreasing EFF, or both. Again, our experiments allow the separate identification of the effects of RISK and EFF.

Three studies used designs in which participants played a series of games against randomly changing opponents, with payoffs varying across games and feedback at the end of each game. Vlaev and Chater (2006) varied the K-index across games and found that the cooperation rate increased with the K-index. Schmidt et al. (2001) and Ahn et al. (2001) examined the impact of variations in 'greed' $(\frac{T-R}{T-S})$ and 'fear' $(\frac{P-S}{T-S})$ on cooperation. These two studies are closely related to our own as greed and fear are alternative measures of temptation and risk (based on a different normalization to those used in the TEMPT and RISK indices). Schmidt et al. (2001) varied the values of R and P across six games while keeping the values of T and S constant and found similar effect sizes of greed and fear on cooperation. Note, however, that an increase in greed could reflect higher temptation or lower efficiency (i.e., TEMPT increases and EFF decreases with greed when T and S are held constant). Similarly, an increase in fear could likewise reflect either an increase in risk or a decrease in efficiency. Ahn et al. (2001) is more closely related to us as they varied the payoffs across four games by using high and low values of T and S but holding R and P constant. Thus, efficiency is kept constant in their study and variation in T and S results in separate variation in RISK and TEMPT. Ahn et al. (2001) found that greed (or TEMPT) has a greater impact than fear (or RISK) on cooperation. Note that all three studies provided feedback between games during the experiment, and therefore cooperation might be affected by the outcome of previous games as well as by payoff changes. Indeed, all three studies report significant feedback effects. In our experiments, no feedback between games is provided.

Finally, Au et al. (2012) and Ng and Au (2016) study the relative risk of cooperation (henceforth riskiness) which they define as $(\frac{R-S}{(R-S)+(T-P)})$, and examine how riskiness and participants' risk attitudes affect cooperation. Au et al. (2012) employed 18, 16, and 28 PDs in three experiments, while Ng and Au (2016) used 24 PDs. No feedback was provided until the end of the experiment in either study. Both studies found that the effect of riskiness of PDs depends on participants' risk attitude: risk-averse participants are more likely to cooperate in a less risky game, while risk-seeking participants are more likely to cooperate in a riskier game. However, the measure of riskiness does not disentangle risk, temptation, and efficiency: riskiness increases as T decreases or R increases. Therefore, increasing cooperation of risk-seeking participants with increasing riskiness might be caused by either decreasing temptation or increasing efficiency or both. The orthogonal variation of payoffs in our Studies 1 and 2 avoids these problems.

2.2. Mengel's meta-analysis

A particularly relevant study for our purposes is Mengel (2018) which examines the relative effect of RISK and TEMPT using data from previously published research supplemented by additional experiments that she conducted either in the lab or on AMT. For the 73 games that were played either as one-shot games or in a random matching protocol, Mengel finds that RISK best explains the variation in cooperation rates, while TEMPT cannot explain this variation after controlling for RISK and EFF.

We report a re-analysis of this dataset, using the same OLS regression specification, in Table 2. The dependent variable is the average cooperation rate. Column 1 reproduces the results reported in Table 3 Column 1 of Mengel (2018). RISK is significantly negatively, and EFF is significantly positively, associated with the average cooperation rate. The coefficient on TEMPT is virtually zero and insignificant.

Table 2. Average cooperation rate regressed on payoff indices using Mengel's (2018)

| | | dataset. | | |
|------------|-------------|---------------|---------------|---------------------|
| | (1) | (2) | (3) | (4) |
| | Full sample | Imposing | Imposing | Imposing |
| | | T > R > P > S | T > R > P > S | T > R > P > S |
| | | | & 2R > T + S | $\& \ 2R \le T + S$ |
| DICK | -0.255*** | -0.142* | -0.045 | -0.178 |
| RISK | (0.061) | (0.074) | (0.123) | (0.105) |
| TEMPT | 0.003 | 0.050 | -0.492 | -0.165 |
| | (0.080) | (0.084) | (0.305) | (0.179) |
| DDD | 0.291*** | 0.360*** | 0.301* | 0.443*** |
| EFF | (0.089) | (0.096) | (0.149) | (0.122) |
| Constant | 0.370*** | 0.218** | 0.304** | 0.370* |
| Constant | (0.084) | (0.097) | (0.130) | (0.200) |
| Adj. R^2 | 0.35 | 0.24 | 0.17 | 0.42 |
| Obs. | 73 | 66 | 36 | 30 |

Notes: Coefficients of OLS models with standard errors in parentheses. * p < 0.1; *** p < 0.05; **** p < 0.01. Model 1 reproduces the estimates of Table 3, Column 1, in Mengel (2018).

In some of the games in the full sample P = S and so defecting is only weakly dominant, while in two games T > R > S > P so the game has two strict equilibria. In Column 2, we restrict the sample to games that meet the first PD condition (i.e., T > R > P > S). The effect of RISK on cooperation substantially decreases and becomes only marginally significant.

Column 3 further restricts the subsample to games that meet both PD conditions (i.e., T > R > P > S and 2R > T + S) and shows that neither RISK nor TEMPT are significantly associated with the variation in average cooperation rates, with the caveat that the sample size is considerably reduced when we restrict attention to games that meet both PD conditions. For comparison we include Column 4 which is based on games meeting the first PD condition but not the second (i.e., T > R > P > S and $2R \le T + S$). This is an attempt to establish whether the reduced effect of RISK in Column 3 compared to Column 2 is due to a strong association between cooperation and RISK when $2R \le T + S$, or whether it reflects low power due to the reduced number of observations. In Column 4 the coefficient on RISK is approximately four times that of Column 3, and although insignificant this suggests that the reduced effect of RISK in Column 3 is driven by excluding games where $2R \le T + S$ where there is a strong association of cooperation with RISK.³

It is important to note that the studies included in Mengel's dataset had their own idiosyncratic reasons for selecting their parameters and the variation between the parameterizations is therefore not entirely systematic. In our experiments we design the payoffs explicitly for comparing the effects of payoff indices. Furthermore, the experiments in Mengel's dataset used different instructional materials and framing of tasks: these differences unrelated to payoffs may affect cooperation across experiments. In our experiments, we control these non-payoff factors by holding them constant within our design.

2.3. A preliminary experiment

We conducted our preliminary experiment with 62 participants playing 15 PD games that meet the two standard PD conditions and vary the RISK, TEMPT and EFF indices over a wide range (see Online Appendix A for the instructions and Online Appendix B for the experimental design details, game parameters, procedures, and additional results). We chose convenient nonnegative payoff parameters to vary the RISK, TEMPT and EFF indices over a wide range yielding a low, medium, and high level for each index similar to the studies that entered Mengel's (2018) dataset.

Across the 15 games, cooperation rates varied between 0.37 and 0.77. In Table 3, we report the effect of payoff indices on cooperation using a linear probability model with participant

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³ One can speculate about why RISK has a strong association with cooperation in games with T+S>2R. It may be that when T+S>2R cooperation increases when RISK is lower (S is higher) because the asymmetric outcome is more appealing for efficiency reasons (as we will see below, EFF is an important consideration). The difficulty of interpreting RISK and EFF when the asymmetric outcome maximizes the sum of payoffs underscores our focus on games where efficiency requires mutual cooperation.

random effects. Robust standard errors are clustered on participants. Using a random effects model allows us to estimate the effects of individual characteristics (i.e., age, gender, nationality, major, spending, and political attitude). The dependent variable is a cooperation dummy, and the explanatory variables are payoff indices (RISK, TEMPT, EFF), with controls for individual characteristics and the round in which the respective game was played.

We find a positive and highly significant coefficient of EFF, whereas neither RISK nor TEMPT have a statistically significant effect on cooperation. An increase in EFF of 0.1 is associated with a 3.99 percentage points higher probability of cooperating. The full model results and robustness checks are in Online Appendix B, Table B3.

Table 3. Determinants of cooperative choice in the preliminary experiment (15 PD games).

| Dependent variable: cooperation dummy | 1 |
|---------------------------------------|---------------------|
| RISK | -0.044 (0.036) |
| ТЕМРТ | -0.083 (0.087) |
| EFF | 0.399*** (0.060) |
| Control variables | Yes |
| Constant | 0.249 (0.340) |
| Within R ² | 0.10 |
| Obs. (Clusters) | 930 (62) |

Notes: Coefficients of a random effects linear probability model with robust standard errors clustered on participants in parentheses. The control variables are round, age, gender, nationality, Business/Economics major, spending, and political attitude. The full results are in Online Appendix B, Table B3. * p < 0.1; *** p < 0.05; **** p < 0.01.

In our preliminary experiment, EFF emerged as the sole payoff index that has a positive and highly significant influence on cooperation. Although the 15 games included in the experiment managed to achieve a large variation in the payoff indices comparable to the studies that entered Mengel's (2018) dataset, this design has the drawback that the induced variation in payoff indices is not fully orthogonal. That is, it gives limited ability to conduct clean non-parametric tests of whether cooperation varies when one index is varied, holding other indices constant. Also, we did not elicit beliefs and so it does not allow us to examine, or control for,

the effect of beliefs on choices. Beliefs are interesting because related research in public goods experiments shows that beliefs strongly influence cooperation (e.g., Frey and Meier (2004); Croson (2007); Fischbacher and Gächter (2010); Dufwenberg et al. (2011)). Game parameters in public goods games do causally shift beliefs and cooperation and because many people are conditional cooperators, increased beliefs increase cooperation (e.g., Gächter and Marino-Fages (2023)).

3. Methods

3.1. Experimental design and procedures for within-subjects Study 1

For Study 1, we create different PDs by varying RISK, TEMPT and EFF orthogonally. This allows us to identify the effect of a single payoff index on behavior while holding constant the other two. First, we fix a *low* and *high* level for each of the three payoff indices. We then generated $2^3 = 8$ payoff matrices representing all possible variations of the two levels across the three payoff indices. The payoffs are presented in Table 4. R = 500 is constant across all PDs, while our experiment has two distinct values of $T \in \{600, 800\}$ and $P \in \{200, 400\}$, and four distinct values of $S \in \{20, 90, 40, 180\}$. This procedure yields the values 0.55 and 0.90 for RISK, 0.17 and 0.38 for TEMPT, and 0.20 and 0.60 for EFF.

Table 4. Payoff parameters for Studies 1 and 2.

| Game | T | R | P | S | RISK | TEMPT | EFF | Mean cooperation rates | | |
|------|-----|-----|-----|-----|------|-------|------|------------------------|---------|------|
| | | | | | | | | Stu | Study 1 | |
| | | | | | | | | UoN | AMT | AMT |
| G1 | 600 | 500 | 200 | 90 | 0.55 | 0.17 | 0.60 | 0.49 | 0.59 | 0.61 |
| G2 | 600 | 500 | 200 | 20 | 0.90 | 0.17 | 0.60 | 0.45 | 0.60 | 0.64 |
| G3 | 800 | 500 | 200 | 90 | 0.55 | 0.38 | 0.60 | 0.36 | 0.47 | 0.59 |
| G4 | 800 | 500 | 200 | 20 | 0.90 | 0.38 | 0.60 | 0.38 | 0.40 | 0.53 |
| G5 | 600 | 500 | 400 | 180 | 0.55 | 0.17 | 0.20 | 0.38 | 0.50 | 0.47 |
| G6 | 600 | 500 | 400 | 40 | 0.90 | 0.17 | 0.20 | 0.33 | 0.48 | 0.50 |
| G7 | 800 | 500 | 400 | 180 | 0.55 | 0.38 | 0.20 | 0.28 | 0.45 | 0.56 |
| G8 | 800 | 500 | 400 | 40 | 0.90 | 0.38 | 0.20 | 0.28 | 0.42 | 0.53 |

Notes: Payoffs in experimental currency.

After reading the instructions (see Online Appendix A), participants completed two tasks presented on the same screen for each PD. First, they indicated their decision (cooperate or

defect) with decisions neutrally labelled as options 'A' and 'B'. The labels were presented in a random order with randomization at the pair level to control for potential presentation effects (i.e., 'A' was the cooperative decision in some games but not in others).

Second, participants indicated their belief about the other person's decision by selecting the likelihood (between 0 and 100 percent) of the other player choosing option 'A'. We did not incentivize belief elicitation to avoid a potential hedging problem (Blanco et al. (2010)) that may occur when both choice task and belief elicitation are incentivized.⁴

To control for potential order effects, we randomized at the pair level the sequence in which the decision and belief elicitation tasks were displayed. To ensure that participants recognize the payoff changes and fully understand how all potential outcomes depend on decisions, participants had to answer eight game-specific control questions about how decisions affect own and other payoff. These questions had to be correctly answered before decisions and beliefs could be entered.

Participants did not receive feedback on the others' choices or the game outcomes until the end of the session. Once participants completed the tasks for all games, we asked them to complete a short post-experimental questionnaire. At the end of the session, one game was randomly chosen, at the pair level, for payment. Participants were reminded of their decisions and informed about the outcome for the randomly chosen game.

We ran our experiments online with two subject pools: students recruited from a volunteer database at the University of Nottingham (UoN, n = 162) and workers recruited on Amazon Mechanical Turk (AMT, n = 160). We did this because students are the typical subject pool for the experiments on PDs which inspired our study (see Section 2) and well suited for studying conceptual questions (see Gächter (2010)). However, given that students tend to be less cooperative than older people (e.g., Arechar et al. (2018); Gächter and Herrmann (2011); List (2004); Matsumoto et al. (2016)), the question of generalizability of results arises: How robust are results on payoff variation for cooperation across subject pools with likely different levels of baseline cooperativeness?

We ran our experiments using the same software (LIONESS Lab, Giamattei et al. (2020)) and near-identical instructions for both subject pools. Because Study 1 was conducted online in both subject pools, we expected a non-negligible attrition rate during gameplay. We used the following procedure to determine payoffs considering potential dropouts. If both

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⁴ Another possibility would be to incentivize either the choice task or belief elicitation. This, however, would complicate the instructions making them more difficult to understand. Moreover, Trautmann and van de Kuilen (2015) find that unincentivized and incentivized elicitation perform equally well in terms of accuracy.

participants completed the entire experiment, they were paid according to the outcome of the randomly chosen game. If one of the pair had dropped out during the experiment, the computer randomly selected the payoff-relevant game and randomly selected one of the four monetary outcomes of the chosen game for payment to the remaining participant. We explained this payment scheme in the instructions.

As we implemented real-time matching of participants in Study 1, we were concerned that decreasing attention might lead to prolonged waiting times. We took several measures to retain attention and encourage successful completion of the experiment. Before participants entered the experiment, we told them to avoid distractions during the experiment. In addition, participants who were inactive for more than 30 seconds (i.e., no mouse movement or no keyboard input) got an alert voice message and a blinking text on their browser. If an inactive participant did not respond to the alert message for a further 30 seconds, they were removed from the session so that the remaining participant could complete the experiment. Three participants (2%) recruited from UoN and 39 of participants (24%) recruited via AMT dropped out during the experiment. The relatively high attrition rate amongst participants recruited via AMT is consistent with similar interactive online experiments (Arechar et al. (2018)).

The sessions lasted for approximately 30 minutes, including the completion of a post-experimental questionnaire. Participants were informed of their payment immediately upon completion of the experiment and were paid within 24 hours. Participants recruited at UoN earned on average £4.79 (SD = £2.33); Participants recruited via AMT earned on average \$5.00 (SD = \$2.43), which amounts to an hourly wage of \$10.5 Further descriptive statistics and comparisons of our subject pools are in Online Appendix C.

3.2. Experimental design and procedures for between-subjects Study 2

For Study 2, we adapt the experimental design of Study 1 to a *between-subjects* design using the eight games of Study 1. The only difference from Study 1 is that each participant plays only *one* one-shot game randomly selected from G1 to G8 shown Table 4. This experiment was pre-registered (AEARCTR-0009784).⁶ The instructions were the same as for Study 1, except for the adaptation to one game play (see Instructions for Study 2 in Online Appendix A).

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⁵ The hourly wage of \$10 compares well to the federal minimum wage \$7.25 at the time of the experiment. The results of Kocher et al. (2008) (in lab public goods games) and Amir et al. (2012) (in AMT public goods games and trust games) suggests that results are robust to higher stakes.

⁶ For the details of preregistration, see https://www.socialscienceregistry.org/trials/9784.

Based on a power calculation we aimed at recruiting 200 participants per game, that is, a total of 1,600 participants.⁷ Because these numbers are infeasible in the UoN laboratory and because our results from Study 1 are largely similar between UoN and AMT anyway (apart from higher baseline levels of cooperation in AMT – see Fig. 1) we ran Study 2 on AMT only.

1,609 participants completed the experiment. The sessions lasted for approximately 15 minutes, including the completion of a post-experimental questionnaire. Participants were informed of their payment immediately upon completion of the experiment. Participants earned on average \$3.13 (SD = \$0.92). Online Appendix C includes the full descriptive statistics.

4. Results

4.1. Results from the within-subject experiment (Study 1)

Cooperation. Across the eight games, cooperation rates vary from 0.28 to 0.49 in UoN and from 0.40 to 0.60 in AMT (see Table 4). On average, UoN participants cooperated in 2.96 of the 8 games, which is significantly lower than AMT participants who cooperated in 3.91 games (Mann-Whitney Z = -2.86, p = 0.004). This is consistent with previous studies, discussed above, that find lower levels of cooperative behavior across student than non-student subject pools. 67% of UoN participants (70% of AMT participants) were switchers, 25% (17%) always defected and 8% (13%) always cooperated.

The left panel of Figure 1 illustrates the average cooperation rates in Study 1 in each of the eight PDs separately by payoff index and sample. Panels (a) and (d) show games connected by a line which only differ in their level of RISK. The line pattern illustrates the Bonferroniadjusted significance levels of non-parametric McNemar tests. We find no significant differences in cooperation rates across low- and high-RISK games for any of the four possible pair-wise comparisons possible in either sample.

Panels (b) and (e) show games that differ only in their level of TEMPT connected by a line. For the UoN sample, we find a significantly lower cooperation frequency as TEMPT increases for two of the four comparisons possible. Similarly, the AMT sample includes one highly significant decrease in the cooperation rates as TEMPT increases. Finally, Panels (c), and (f) show games that differ only in their level of EFF connected by a line. The UoN sample

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⁷ In Study 1, TEMPT emerged as the most important of the three indices in explaining cooperation. The cooperation rate under low TEMPT was 0.4 vs 0.6 under high TEMPT, which turned out to be the biggest effect size. Given this treatment difference, a 5% significance test of the equality of two proportions would have 95% power with a sample size of 160 per treatment. To account for heterogeneity on AMT, we planned to recruit 200 participants for each of the 8 games.

provides strong evidence for a positive effect of EFF on cooperation as we find that three out of four comparisons show at least a weakly significant increase in the cooperation frequency as EFF increases. The AMT sample shows one weakly significant increase in the cooperation frequency as EFF increases. We will complement these results with a regression analysis reported below, but before we do so, we discuss how payoffs affect beliefs.

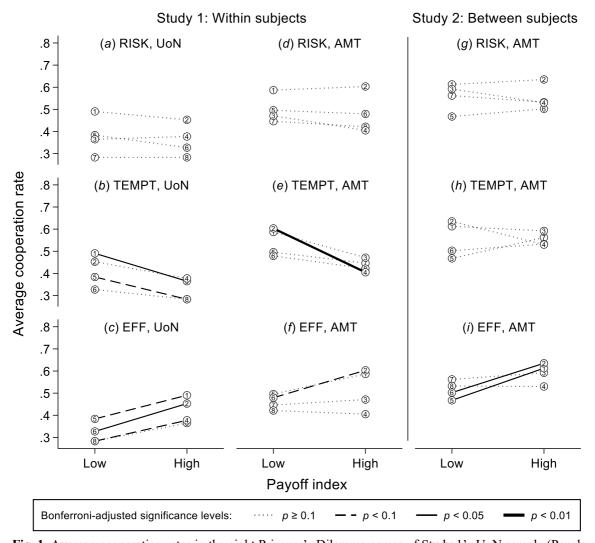


Fig. 1. Average cooperation rates in the eight Prisoner's Dilemma games of Study 1's UoN sample (Panels a-c), Study 1's AMT sample (Panels d-f) and Study 2 (Panels g-i). The line patterns indicate the Bonferroni-adjusted significance levels of two-sided McNemar's tests (Study 1) and Fisher's exact tests (Study 2). The game number is shown in the respective marker. See Online Appendix D, Table D1-D2 for the uncorrected p-values.

Beliefs. As beliefs have been identified as an important driver of cooperative behavior in similar games, such as the public good game (e.g., Croson (2007); Fischbacher and Gächter (2010); Gächter and Renner (2018)) and the sequential prisoner's dilemma game (e.g., Baader et al. (2022)), we now examine how the variation in payoffs affects beliefs. Figure 2 shows the average expected likelihood that the other player cooperates separately by payoff index and

sample. On average, AMT participants held higher average cooperative beliefs than UoN participants (Mann-Whitney Z = -2.44, p = 0.015) but in terms of belief accuracy (average belief compared to cooperation rate) we find a weakly significantly higher accuracy in the UoN subject pool (for details see Online Appendix E, Table E1).

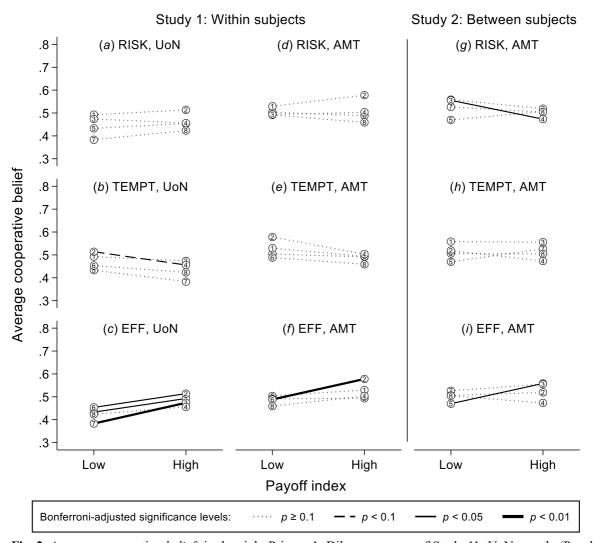


Fig. 2. Average cooperative *beliefs* in the eight Prisoner's Dilemma games of Study 1's UoN sample (Panels *a-c*), Study 1's AMT sample (Panels *d-f*) and Study 2 (Panels *g-i*). The line patterns indicate the Bonferroni-adjusted significance levels of two-sided Wilcoxon signed-rank tests (Study 1) and Mann-Whitney tests (Study 2). The game number is shown in the respective marker. See Online Appendix D, Table D3-D4 for the uncorrected *p*-values.

In Panels (a) and (d) games that differ only in their level of RISK, but not in TEMPT or EFF, are connected by a line. Beliefs across these two games are directly comparable. No clear effect of a change in RISK on average beliefs emerges, as average beliefs decrease in some games but increase in others. A series of non-parametric Wilcoxon signed-rank tests shows insignificant differences in the average beliefs in both the UoN and the AMT sample. Panels (b)

and (e) illustrate pairs of games that only differ in TEMPT. Beliefs about the other player's cooperativeness decrease as TEMPT increases, but the effect is only marginally significant for one of the four game pairs in the UoN sample. Panels (c) and (f) show the pairs of games differing in EFF only. We find that an increase in EFF is associated with an increase in the average cooperative belief for almost all pairs of games. The difference between the low- and high-EFF games is highly significant for one game pair and significant for two of the game pairs in the UoN sample. For the AMT sample, we find highly significant differences for one of the four game pairs. The next step in our analysis is a regression analysis that controls for beliefs.

Regression results. In Table 5, we report the effect of payoff indices on cooperation and belief using linear (probability) models with participant random effects and robust standard errors clustered on participants separately for both samples. In all models, we control for the subject pool, individual characteristics, and task characteristics (i.e., the round in which the respective game was played, whether the decision task or belief task appeared at the top of the screen and labelling of cooperative choice as A or B). The full model results are in Online Appendix F, Table F1.

The models in Columns 1-2 show that that the effect of RISK on cooperation is small in magnitude and insignificant in both samples. TEMPT appears to be the most influential determinant of cooperation. The coefficients on TEMPT are negative, highly significant, and show a larger effect than EFF and RISK in both samples. An increase in TEMPT of 0.1 is associated with a 4.08 (5.06) percentage points lower probability of cooperating in the UoN (AMT) sample. EFF also appears as an influential determinant of cooperation (although the effect size is smaller than TEMPT). A 0.1 increase in EFF increases cooperation by 2.45 percentage points for UoN participants and 1.45 percentage points for AMT participants.⁸

In Columns 3-4, we estimate the effect of payoff indices on beliefs, which is an important co-variate of our behavioral outcome measure (UoN: $r_s = 0.48$, p < 0.001; AMT: $r_s = 0.41$, $p_s < 0.001$; Pooled samples). We find a significantly negative effect of TEMPT on belief across both samples. EFF positively affects beliefs in both samples, with a highly significant coefficient of EFF for UoN and a weakly significant and smaller coefficient for AMT.

there was a differential effect of RISK and/or TEMPT across high versus low EFF games. For the AMT sample we find that the effect of TEMPT is stronger in the high EFF games. We find no differential effect of TEMPT in the UoN sample and no differential effect of RISK in either sample. See Online Appendix F, Table F3 for details.

⁸ We also ran the regressions including a high EFF dummy interacted with RISK and TEMPT to examine whether there was a differential effect of RISK and/or TEMPT across high versus law EFF games. For the AMT sample

Table 5. Payoff indices, beliefs, and cooperation in Studies 1 and 2.

| | | With | in-subjects ex | periment (Stu | ıdy 1) | | Between-sub | jects experin | nent (Study 2) |
|-------------------------|-------------|-------------|----------------|---------------|---------------------|---------------------|-------------|---------------|---------------------|
| Dependent variable: | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
| | UoN | AMT | UoN | AMT | UoN | AMT | AMT | AMT | AMT |
| | Cooperation | Cooperation | Belief | Belief | Cooperation | Cooperation | Cooperation | Belief | Cooperation |
| RISK | -0.094 | -0.073 | 0.001 | -0.014 | -0.094 | -0.066 | -0.012 | -0.042 | 0.012 |
| | (0.062) | (0.062) | (0.038) | (0.051) | (0.058) | (0.062) | (0.067) | (0.030) | (0.065) |
| TEMPT | -0.408*** | -0.506*** | -0.147** | -0.174** | -0.323*** | -0.431*** | -0.032 | -0.029 | -0.015 |
| | (0.108) | (0.117) | (0.062) | (0.078) | (0.107) | (0.112) | (0.113) | (0.050) | (0.110) |
| EFF | 0.245*** | 0.145** | 0.155*** | 0.076* | 0.155*** | 0.113 | 0.179*** | 0.048* | 0.152*** |
| | (0.058) | (0.073) | (0.033) | (0.042) | (0.059) | (0.072) | (0.059) | (0.026) | (0.057) |
| Belief | | | | | 0.582*** (0.061) | 0.434*** (0.068) | | | 0.565*** (0.053) |
| Control variables | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | 0.180 | 0.575*** | 0.476*** | 0.552*** | -0.098 | 0.334** | 0.588*** | 0.771*** | 0.153 |
| | (0.193) | (0.155) | (0.098) | (0.105) | (0.164) | (0.139) | (0.094) | (0.047) | (0.097) |
| (Within) R ² | 0.06 | 0.04 | 0.07 | 0.13 | 0.15 | 0.06 | 0.11 | 0.64 | 0.17 |
| Obs. (Clusters) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) | 1601 | 1601 | 1601 |

Notes: Coefficients of a random effects linear probability model (Cols. 1-2, 5-6) or linear model (Cols. 3-4) with robust standard errors clustered on participants in parentheses. Coefficients from a linear probability model (Cols. 7, 9) or linear model (Col. 8) with robust standard errors in parentheses. The control variables are round, order of tasks, order of choices, age, gender, ethnicity, Business/Economics major (UoN only), spending/income, political attitude, and previous experience in experiments. Full estimation results are in Online Appendix F, Table F1-F2. * p < 0.1; *** p < 0.05; **** p < 0.01.

The positive correlation of beliefs and cooperation is a common result in the literature on related social dilemma games (e.g., Dufwenberg et al. (2011)). It is consistent with experiments that causally manipulated beliefs (e.g., Frey and Meier (2004)) or held beliefs constant via the strategy method (e.g., Fischbacher and Gächter (2010); Gächter et al. (2022)). On the aggregate level, this can be taken as evidence for conditional cooperation, although this masks a substantial individual-level heterogeneity on the correlation between individual beliefs and behavior (see Online Appendix G for an illustration and discussion).

Columns 5-6 present results from the model which includes the payoff indices and beliefs as explanatory variables. For both samples, the coefficient of TEMPT and EFF are reduced in size when Belief is added to the model. For AMT, the effect of EFF even becomes statistically insignificant. This implies that the total effect of these two payoff indices on cooperation is (partially) mediated through beliefs.

More formally, we can decompose the total effect of the payoff indices into direct and indirect components via the mediator variable Belief using the method proposed by Baron and Kenny (1986). See Online Appendix H for the details. For example, the total effect of TEMPT on cooperation in the UoN sample, comprising direct and indirect effects, is given by the highly significant and negative coefficient in Column 1 (b = -0.408, p < 0.001). The Baron and Kenny method proposes that the indirect effect through Belief can be approximated by multiplying the direct effect of TEMPT on the mediator Belief (b = -0.147, p = 0.017; Column 3) with the direct effect of the mediator Belief on cooperation (b = 0.582, p < 0.001; Column 5), yielding a significant negative indirect effect (b = -0.086, p = 0.011), which accounts for 21% of the total effect in the UoN sample. For the AMT sample, we also find a significant indirect effect of TEMPT mediated through Belief (b = -0.076, p = 0.018), which accounts for 15% for the total effect. Regarding the indirect effect of EFF mediated through Belief on behavior, we find a (highly) significant indirect effect in both samples (UoN: b = 0.090, p < 0.001; AMT: b = 0.033, p = 0.043). In UoN the indirect effect accounts for 37% of the total effect and in AMT it accounts for 23%.

All regressions include task characteristics and individual characteristics as controls. The individual characteristics are generally insignificant (see Online Appendix F, Table F1 for details). Round is significantly negative (except in model (6)) despite no feedback between

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⁹ While this is a frequently used methodology, it is important to acknowledge that it rests on relatively strong assumptions of linear models and the absence of confounding effects between the mediator and outcome variable (for a discussion of mediation analysis in economics, see for example, Celli (2022)).

games. This is consistent with "virtual learning" (Weber (2003)) that has also been observed in public goods games (e.g., Neugebauer et al. (2009)).

As a final step in our analysis of Study 1, we take advantage of the within-subject nature of the data and examine the consistency of cooperative behavior across the two different levels of a payoff index. We evaluate consistency using an assumption of (weak) monotonicity: for instance, someone who cooperates under high TEMPT should cooperate under low TEMPT. To count the number of violations in monotonicity, we compare twelve pairs of games: 4 pairs which only differ in RISK, 4 pairs which only differ in TEMPT, and 4 pairs which only differ in EFF. For RISK (TEMPT), cooperating in a higher RISK (TEMPT) game but defecting in a lower RISK (TEMPT) game holding other payoffs indices constant is counted as a violation of monotonicity. For EFF, cooperating in a low EFF game but defecting in a high EFF game holding other payoff indices constant is counted as a violation.

Figure I1 in Online Appendix I shows the number of violations by each payoff index. In both UoN and AMT samples, participants violate monotonicity assumptions at least once at the following rates:

- RISK: 37% (43%) of UoN (AMT) participants.
- TEMPT: 31% (32%) of UoN (AMT) participants.
- EFF: 30% (36%) of UoN (AMT) participants.

There are no systematic subject pool differences in the degree of monotonicity violations for any of the payoff indices (Fisher's exact tests, all $p \ge 0.304$). These results imply that in the UoN sample the findings of Figure 1 and Table 5 are not due to systematic and robust indexspecific inconsistencies in behavior. In the AMT sample the higher rate of non-monotonic choices in RISK compared to TEMPT and EFF might, however, have contributed to insignificant results in RISK.

4.2. Discussion of Study 1

RISK does not have a significant impact on cooperation in Study 1. In contrast, TEMPT has a highly significant negative effect on cooperation and EFF has a significant positive effect on cooperation. In addition, we find similar effects of TEMPT and EFF on beliefs. Beliefs appear as a partial mediator of the effect of TEMPT and EFF, accounting for a substantial share of the total effect. These results are observed in both subject pools, except for the significant effect of EFF on cooperation in the AMT sample disappearing after controlling for beliefs.

One concern about these results is that they may be sensitive to the within-subject design nature of Study 1. The within-subject payoff variations might have allowed participants to learn through enhanced experience in game play or induced participants to change their decisions either because of a perceived experimenter demand effect ("payoffs changed, so I should change my decisions too") or because changing payoffs makes them more salient (for a discussion of within- vs. between-subjects designs see Charness et al. (2012)). To address these issues, we designed Study 2 where participants played only one game, and games varied between subjects.

4.3. Results from the between-subjects experiment (Study 2)

Results on cooperation. Across the eight games, cooperation rates vary from 0.47 to 0.64. Figure 1 Panels (g)-(i) illustrate the average cooperation rates in each of the eight PDs by payoff index. We find no significant differences in cooperation rates across low- and high-RISK games for any of the four possible pair-wise comparisons. The same is true when comparing low- and high-TEMPT games. We do find significant differences in cooperation rates across low- and high-EFF games for two pair-wise comparisons. In both pair-wise comparisons, cooperation rates increase as EFF increases.

Results on beliefs. Figure 2 Panels (g)-(i) shows the average expected likelihood that the other player chooses 'cooperate' separately for each payoff index. Beliefs about other player's cooperativeness decrease as RISK increases: the effect is significant at the 5% level for one of the four pairs. TEMPT has an ambiguous effect on beliefs as we observe an unclear pattern between beliefs and TEMPT. Increasing EFF strengthens the beliefs about other's cooperativeness: the effect of EFF is significant at the 5% level for one of the four pairs.

Regression results. Next, in Table 5 Columns 7-9, we report the effect of payoff indices using linear (probability) models. Again, we focus on payoff indices RISK, TEMPT, and EFF, and relegate the full regression results to Online Appendix F, Table F2. The analysis parallels Study 1 but is adapted to the strict one-shot nature of the data.

Column 7 reveals a positive and highly significant effect of EFF on cooperation. The coefficients for RISK and TEMPT are not significantly different from zero. Similarly, Column 8 indicates a positive and weakly significant effect of EFF on Belief, while RISK and TEMPT are not significantly different from zero. ¹⁰ Cooperation and beliefs again appear highly

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 $^{^{10}}$ The relatively high R^2 in this regression model is particularly noteworthy. Online Appendix F, Table F2 reveals that the only highly significant control variable is the labelling of strategies. Taken together, this suggests that most participants in Study 2 expected the other player to choose the strategy that was labelled as A, independent of the variation in payoff parameters.

significantly correlated ($r_s = 0.38$, p < 0.001). Column 9 presents the results of the model that includes the three payoff indices and Belief. The coefficient for EFF is highly significant but smaller compared to that reported in Column 7. The coefficient for Belief is highly significant, positive, and similar in size compared to Study 1. A mediation analysis reveals that the significant total effect of EFF on cooperation comprises a significant indirect effect through Belief (b = 0.028, p = 0.035), which accounts for 15% of the total effect.

4.4. Discussion of Study 2

Study 2 tested the role of payoff parameters across eight one-shot PDs in between-subject experiments (with $n \approx 200$ per game). This avoids potential learning through enhanced experience in game play or experimenter demand effects and saliency effects that come from within-subject variation in payoffs. The Study 2 results suggest that the only robust payoff index effect works through EFF: a higher EFF results in a higher cooperation rate. RISK and TEMPT are both insignificant. In our studies of one-shot PDs, EFF is the only variable that robustly influences cooperation in both within- and between-subject designs.

An important question is how sensitive our results are to the specific indices that we use. In the next section, we analyze three related indices. These indices are normalized loss and normalized gain, which are akin to RISK and TEMPT, and the K-index, which resembles EFF.

5. Related payoff indices: normalized loss, normalized gain, and K-index

In this section, we report evidence on related payoff indices that have received attention in the experimental PD literature. Unlike our payoff indices, these are defined on three or four payoff parameters, as will become clear below.

5.1. Normalized loss and normalized gain

A game's payoff matrix can be normalized by subtracting P from all payoffs of the PD payoff matrix (see Table 1) and dividing by R - P (see, e.g., Stahl (1991); Dal Bó and Fréchette (2018); Embrey et al. (2018)). This yields a payoff for mutual cooperation of 1 and a payoff for mutual defection of 0 in the normalized payoff matrix. Thus, the game's efficiency—defined as the payoff difference between mutual cooperation R and mutual defection P—is set to 1. Normalized loss is given by $\frac{S-P}{R-P} = -l$ and therefore $l = \frac{P-S}{R-P}$, which captures the risk of cooperation against a defector (similar to the RISK index but normalized by R-P instead of P). Normalized gain is defined as $\frac{T-P}{R-P} = 1 + g$ which implies that $g = \frac{1}{R-P}$

 $\frac{T-R}{R-P}$, that is, g measures the gain from defecting against a cooperator (similar to the TEMPT index but normalized by R-P instead of R). Across our eight games, l and g vary orthogonally within the sets of four low/high-EFF games (see Online Appendix J for a summary and illustration).

Table 6 reports regression results focusing on the normalized payoff indices for *low-EFF* games, with the full results including controls reported in Online Appendix J, Table J2-J3. In Study 1, l has no statistically significant effect on cooperation in either the UoN or AMT samples, while g has a highly significant negative effect on cooperation in the UoN sample only (Col. 1-2). Similarly, we do not find a statistically significant effect of l on Belief for either sample and we find a significant negative effect of g on Belief in UoN only (Col. 3-4). The model that includes Belief as an explanatory variable reveals weakly significant negative effects of l and g on cooperation for UoN only, and a highly significant positive effect of Belief for both samples (Col. 5-6). The total effect of g on cooperation shown in Column 1 comprises a 33% indirect effect mediated through Belief which is negative and highly statistically significant (b = -0.013, p = 0.009). In Study 2, we find no significant effects of l or l on cooperation or beliefs (Col. 7-8). However, the full model including Belief as an explanatory factor shows a highly significant positive effect of Belief on cooperation (Col. 9).

Table 7 reports regression results for the *high-EFF* games (see Online Appendix J, Table J4-J5 for the full results). In Study 1, l has no statistically significant effect and g has a highly significant negative effect on cooperation across both samples (Col. 1-2). Again, we do not find a statistically significant effect of l on belief for either the UoN or AMT sample. g has a significant negative effect on belief in the AMT sample only (Col. 3-4). When estimating the model which includes belief as an explanatory variable, we find no statistically significant effects of l, but we do find highly significant negative effects of g on cooperation in the UoN and AMT samples. The coefficient for belief is highly significant and positive for both samples (Col. 5-6). The total effect of g on cooperation shown in Col. 1-2 can be decomposed in direct and indirect effects. For the UoN sample, the indirect effect mediated through beliefs is statistically insignificant (b = -0.019, p = 0.158). Yet for AMT, the total effect comprises a significant and negative 14% indirect effect mediated through Belief (b = -0.034, p = 0.019). In Study 2, we only find a weakly significant negative effect of g on cooperation and no significant effect of the normalized payoff indices on Belief (Col. 7-8). The full model, which includes Belief, shows no significant effects of l or g but a highly significant positive effect of

Belief on cooperation (Col. 9). We find no evidence for a significant indirect effect of g on cooperation mediated through Belief (b = -0.007, p = 0.300).

Additional evidence for the role of normalized loss and gain on cooperation can be obtained from our preliminary experiment (see Online Appendix B, Fig. B2 and Table B4). Our regression analysis reveals highly significant negative effects of both, normalized loss and gain on cooperation (b = -0.011, p < 001; b = -0.012, p < 001; resp.). Note, however, that the 15 PD games included in the pre-test do not provide an orthogonal variation in normalized loss and gain.

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¹¹ An alternative way to jointly test the effect of normalized indices and beliefs on behavior is to create a composite index of these factors. Online Appendix J, Table J6 shows that cooperation behavior is jointly affected by the games' incentives as captured by the normalized indices and expected behavior in others.

Table 6. Normalized loss l, normalized gain g, beliefs, and cooperation in low-EFF games.

| | Within-subjects experiment (Study 1) | | | | | | | | nent (Study 2) |
|--------------------------|---------------------------------------|---------------------------------------|----------------------------------|----------------------------------|---------------------------------------|---------------------------------------|-------------------------------|--------------------------|-------------------------------|
| Dependent variable: | (1) Study 1: UoN Cooperation | (2) Study 1: AMT Cooperation | (3) Study 1: UoN Belief | (4) Study 1: AMT Belief | (5) Study 1: UoN Cooperation | (6) Study 1: AMT Cooperation | (7) Study 2 Cooperation | (8) Study 2 Belief | (9) Study 2 Cooperation |
| Normalized loss <i>l</i> | -0.030 (0.020) | -0.011 (0.022) | 0.008 (0.013) | -0.017 (0.018) | -0.034* (0.019) | -0.002 (0.023) | -0.004 (0.024) | -0.003 (0.010) | -0.002 (0.023) |
| Normalized gain g | -0.039*** (0.014) | -0.025 (0.017) | -0.023** (0.009) | -0.010 (0.012) | -0.026* (0.014) | -0.020 (0.016) | 0.021 (0.017) | -0.002 (0.007) | 0.022 (0.016) |
| Belief | | | | | 0.571*** (0.082) | 0.529*** (0.083) | | | 0.512*** (0.077) |
| Control variables | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | 0.178 (0.223) | 0.345** (0.173) | 0.464*** (0.112) | 0.499*** (0.121) | -0.090 (0.190) | 0.074 (0.146) | 0.536*** (0.128) | 0.753*** (0.068) | 0.150 (0.133) |
| (Within) R ² | 0.05 | 0.01 | 0.08 | 0.10 | 0.12 | 0.04 | 0.12 | 0.64 | 0.16 |
| Obs. (Clusters) | 616 (154) | 476 (119) | 616 (154) | 476 (119) | 616 (154) | 476 (119) | 802 | 802 | 802 |

Notes: Coefficients of a random effects linear probability model (Cols. 1-2, 5-6) or linear model (Cols. 3-4) with robust standard errors clustered on participants in parentheses. Coefficients from a linear probability model (Cols. 7, 9) or linear model (Col. 8) with robust standard errors in parentheses. The control variables are round, order of tasks, order of choices, age, gender, ethnicity, Business/Economics major (UoN only), spending/income, political attitude, and previous experience in experiments. Full estimation results are in Online Appendix J, Table J2-J3. * p < 0.1; *** p < 0.05; **** p < 0.01.

Table 7. Normalized loss l, normalized gain g, beliefs, and cooperation in high-EFF games.

| | Within-subjects experiment (Study 1) | | | | | | | | nent (Study 2) |
|--------------------------|---------------------------------------|---------------------------------------|----------------------------------|----------------------------------|---------------------------------------|---------------------------------------|-------------------------------|--------------------------|-------------------------------|
| Dependent variable: | (1) Study 1: UoN Cooperation | (2) Study 1: AMT Cooperation | (3) Study 1: UoN Belief | (4) Study 1: AMT Belief | (5) Study 1: UoN Cooperation | (6) Study 1: AMT Cooperation | (7) Study 2 Cooperation | (8) Study 2 Belief | (9) Study 2 Cooperation |
| Normalized loss <i>l</i> | -0.106 (0.134) | -0.163 (0.140) | -0.046 (0.083) | 0.052 (0.089) | -0.075 (0.123) | -0.184 (0.145) | -0.010 (0.142) | -0.104 (0.064) | 0.052 (0.138) |
| Normalized gain g | -0.133*** (0.047) | -0.237*** (0.056) | -0.029 (0.028) | -0.078** (0.033) | -0.114** (0.047) | -0.204*** (0.054) | -0.083* (0.050) | -0.012 (0.022) | -0.076 (0.048) |
| Belief | | | | | 0.668*** (0.071) | 0.434*** (0.092) | | | 0.604*** (0.075) |
| Control variables | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | 0.354* (0.209) | 0.845*** (0.165) | 0.601*** (0.105) | 0.643*** (0.110) | -0.049 (0.188) | 0.567*** (0.160) | 0.791*** (0.121) | 0.812*** (0.056) | 0.301** (0.132) |
| (Within) R ² | 0.07 | 0.09 | 0.04 | 0.20 | 0.18 | 0.10 | 0.12 | 0.65 | 0.18 |
| Obs. (Clusters) | 616 (154) | 476 (119) | 616 (154) | 476 (119) | 616 (154) | 476 (119) | 799 | 799 | 799 |

Notes: Coefficients of a random effects linear probability model (Cols. 1-2, 5-6) or linear model (Col. 3-4) with robust standard errors clustered on participants in parentheses. Coefficients from a linear probability model (Cols. 7, 9) or linear model (Col. 8) with robust standard errors in parentheses. The control variables are round, order of tasks, order of choices, age, gender, ethnicity, Business/Economics major (UoN only), spending/income, political attitude, and previous experience in experiments. Full estimation results are in Online Appendix J, Table J4-J5. * p < 0.1; *** p < 0.05; **** p < 0.01.

5.2. A summary index of PD parameters: The K-index

Recall that the K-index is defined as $\frac{R-P}{T-S}$ (Rapoport (1967)). It is based on all four PD payoff parameters, and it captures the gains from mutual cooperation over mutual defection, R-P, relative to the range of payoffs, T-S. Because T>R>P>S, the K-index \in (0,1). For the K-index values of our games, see Table J1 in Online Appendix J.

The K-index is an index of cooperation: the higher the K-index, the more beneficial is mutual cooperation, that is, the lower is the conflict of interest between collective benefit and private gain (see also Balliet and Van Lange (2013)). We therefore expect cooperation to increase in the K-index, in line with previous literature (for recent meta-analyses of PD games using the K-index, see, e.g., Balliet and Van Lange (2013); Thielmann et al. (2020); Yuan et al. (2022); and Spadaro et al. (2022)).

The K-index is interesting because it is a summary index of the severity of the cooperation problem. But for our purposes, the K-index analysis that follows below also serves as a robustness check for EFF, which shares the same numerator, R - P, with the K-index.

Fig. 3 shows how the K-index of a game is related to the average cooperation rate, separately for each study and subject pool. In line with the previous literature, we see that the K-index and cooperation are positively related in all studies and all subject pools, although with some interesting differences between them.

- In Studies 1 and 2, the K-index is between 0.13 and 0.59. Interestingly, for all eight levels of the K-index, cooperation rates are higher in the AMT subject pool, where cooperation rates range from 0.40 to 0.60, whereas in the UoN subject pool, they range from 0.28 to 0.49. Cooperation rates are positively correlated with the K-index: this correlation is highly significant for UoN participants, whereas it is marginally insignificant for AMT participants (UoN: $r_s = 0.90$, p = 0.002; AMT: $r_s = 0.62$, p = 0.102).
- In Study 2, which only used AMT participants, cooperation rates range from 0.47 to 0.61, and the correlation of cooperation rates and the K-index is similar to Study 1 for AMT participants: $r_s = 0.59$, p = 0.120.

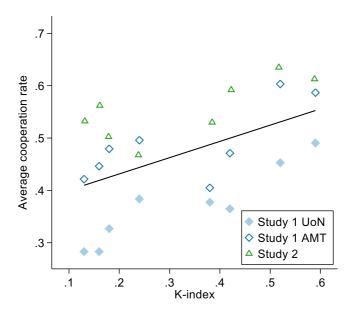


Fig. 3. Average cooperation rates for each level of a game's K-index, by study and subject pool. *Note*: The black line shows the predicted values from a linear regression.

In the pooled dataset, disregarding study, and subject pool, we have 24 distinct average cooperation rates. Here, the Spearman correlation is $r_s = 0.46$, p < 0.023. A simple OLS regression of cooperation rate on K-index returns a coefficient for the K-index of 0.309 (with a 95% CI of [0.081, 0.536]; $R^2 = 0.26$), which is slightly lower than the estimated coefficient (0.44) of the K-index in the meta-analysis of Yuan et al. (2022) (see their Table 3). Thus, overall, a PD's K-index predicts its average cooperation rate well.

Table 8 reports the effect of the K-index on cooperation and beliefs. In Study 1, we find a positive and highly significant effects on cooperation across the UoN and AMT samples (Col. 1-2) as well as positive and highly significant effects on Belief for both samples (Col. 3-4). The model that includes Belief as an explanatory variable reveals positive and highly significant effects of both, the K-index and Belief, for the UoN and AMT samples (Col. 5-6).

Decomposing the total effects of the K-index on cooperation shown in Columns 1-2 reveals that positive and highly significant indirect effects mediated through beliefs account for 30% of the total effect in the UoN sample and 19% of the total effect in the AMT sample (UoN: b = 0.116, p < 0.001; AMT: b = 0.055, p = 0.009).

Similarly, we find a highly significant positive effect of the K-index on cooperation and belief in Study 2 (Col. 7-8). For the full model, including belief, the coefficient for the K-index is highly significant and positive albeit somewhat reduced in size. The coefficient for belief is also positive and highly significant (Col. 9). The total effect of the K-index on belief comprises a 17% positive and significant indirect effect mediated through belief (b = 0.039, p = 0.018).

Additional support for the role of the K-index in explaining cooperation comes from the analysis of our preliminary experiment (see Online Appendix B, Fig. B3 and Table B4), which reveals a highly significant and positive effect of the K-index on cooperation (b = 0.442, p < 001).

5.3. Discussion

In this section we have investigated the robustness of our conclusions by looking at three closely related payoff indices: normalized loss, which resembles RISK; normalized gain, which resembles TEMPT; and the K-index, which resembles EFF. These indices all share the same numerator with our respective indices but have different denominators.

The results based on these alternative payoff indices largely confirm the findings based on RISK, TEMPT and EFF. In neither Study 1 nor 2 do we find a significant effect of normalized loss (akin to RISK) on cooperation. We do however find some evidence that normalized gain matters for cooperation, particularly in games with high efficiency. These results should be interpreted with caution as we did not design the experiments for a controlled variation of the normalized indices and thus the variation in normalized loss and gain is larger in high-efficiency games.

Table 8. K-index, beliefs, and cooperation.

| Within-subjects experiment (Study 1) | | | | | | | Between-sub | jects experin | nent (Study 2) |
|--------------------------------------|-------------|-------------|------------|-----------|-------------|-------------|-------------|---------------|----------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) |
| Dependent | Study 1: | Study 1: | Study 1: | Study 1: | Study 1: | Study 1: | Study 2 | Study 2 | Study 2 |
| variable: | UoN | AMT | UoN | AMT | UoN | AMT | Cooperation | Belief | Cooperation |
| | Cooperation | Cooperation | Belief | Belief | Cooperation | Cooperation | | | |
| IZ : 1 | 0.382*** | 0.297*** | 0.199*** | 0.126*** | 0.265*** | 0.242*** | 0.228*** | 0.069** | 0.189*** |
| K-index | (0.073) | (0.084) | (0.041) | (0.049) | (0.072) | (0.083) | (0.071) | (0.032) | (0.070) |
| D 1' C | | | | | 0.583*** | 0.440*** | | | 0.564*** |
| Belief | | | | | (0.061) | (0.068) | | | (0.053) |
| Control variables | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| | -0.025 | 0.342** | 0.435*** | 0.482*** | -0.279* | 0.127 | 0.570*** | 0.730*** | 0.158** |
| Constant | (0.180) | (0.140) | (0.092) | (0.096) | (0.151) | (0.123) | (0.073) | (0.038) | (0.078) |
| (Within) R ² | 0.06 | 0.03 | 0.07 | 0.13 | 0.14 | 0.06 | 0.11 | 0.64 | 0.17 |
| Obs. (Clusters) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) | 1601 | 1601 | 1601 |

Notes: Coefficients of a random effects linear probability model (Cols. 1-2, 5-6) or linear model (Cols. 3-4) with robust standard errors clustered on participants in parentheses. Coefficients from a linear probability model (Col. 7, 9) or linear model (Col. 8) with robust standard errors in parentheses. The control variables are round, order of tasks, order of choices, age, gender, ethnicity, Business/Economics major (UoN only), spending/income, political attitude, and previous experience in experiments. Full estimation results are in Online Appendix J, Table J7-J8. * p < 0.1; ** p < 0.05; *** p < 0.01.

6. Towards an explanation of our results

How can social preferences explain our results on the positive impact for cooperation of EFF (and the K-index) in all experiments? Why does TEMPT matter in the within-subject study but not in the between-subject study? Why does RISK never matter? A full-fledged formal analysis of what theories of social preferences predict in our games is beyond the scope of this paper. Instead, we use the basic psychological motives incorporated in the various theories and their experimental tests as likely sources of psychological considerations that players of prisoner's dilemma games might entertain. Our answers are inevitably somewhat speculative because we did not set up the experiments to test a particular theory (unlike, e.g., the horserace conducted by Miettinen et al. (2020)). Participants in one-shot games are also unlikely to be performing a full-fledged strategic analysis of the games they play but rather employ heuristics based on the comparative attractiveness of various possible outcomes (see, e.g., the approaches of Stewart et al. (2016) and Lugrin et al. (2024) who use eye-tracking methods in 2x2 games). In the following we discuss considerations that might guide cooperation decisions in our experiments.

In abstract, anonymous games with monetary payoffs, like in ours, a likely consideration for many people is based on their distributional preferences. Many people are *inequality averse* both when it is to their advantage and when it is to their disadvantage (see Fehr and Schmidt (1999) and Bolton and Ockenfels (2000) for the theoretical arguments and supporting evidence, and Blanco et al. (2011) and Beranek et al. (2015) for empirical estimates on inequality aversion parameters). Inequality aversion renders strategy combinations resulting in equal payoffs [(R, R)] and (P, P) somewhat more attractive (and thereby 'focal' or 'salient' for inequality averse people) than strategies resulting in unequal payoffs [(T, S)] and (S, T). Combined with the fact that R > P, inequality aversion and preferring more money over less money makes mutual cooperation attractive, and this attraction increases the larger R - P (and hence EFF) is. Many people's distributional preferences also contain *preferences for efficiency*, whereby people are willing to incur some cost to maximize payoffs (e.g., Charness and Rabin (2002); Engelmann and Strobel (2004)). Therefore, for efficiency-concerned people, cooperation becomes more likely as EFF increases.

The attractiveness of EFF is further reinforced by a range of well-established motives beyond distributional preferences whose relevance likely also increases as EFF increases. The following motives are also likely to positively influence beliefs about the likelihood of cooperation by a player's opponent thereby further increasing the likelihood of cooperation:

- Warm glow, altruism, and Kantian morality, according to which people derive some utility from the act of cooperating (Andreoni (1995); Palfrey and Prisbrey (1997)) and cooperating is "the right thing to do" (e.g., Alger and Weibull (2013));
- *Team reasoning*, the idea that players view their opponent as a team member and play the action that maximizes team payoffs, which prescribes mutual cooperation (e.g., Bacharach (2006));
- Magical thinking, which is a belief that one's own act of cooperation makes cooperation by the opponent more likely (Shafir and Tversky (1992); Daley and Sadowski (2017));
- Reciprocity, guilt aversion, and conditional cooperation by which people are more likely to cooperate if they expect others to cooperate and believe their opponent expects them to cooperate (e.g., Dufwenberg et al. (2011); Fischbacher and Gächter (2010)).

Why does TEMPT matter in the within-subject experiments of Study 1 but not in the between-subjects experiments of Study 2? A candidate behavioral explanation is related to *salience*. A stimulus is salient if it automatically attracts a decision maker's attention and one source of salience is contrast with surroundings (see Bordalo et al. (2022) for a review of the literature). In the within-subject experiments of Study 1 participants played eight games with changing parameters (*T*, *S*, *P* relative to a fixed *R*) thereby creating contrasts that made changes in TEMPT salient, whereas in the between-subjects experiments of Study 2 participants only played one game with a given TEMPT parameter (and hence no contrast due to change). This means that the stimulus of TEMPT attracts more attention, and hence is more salient, in Study 1 than in Study 2. Because TEMPT is an appeal to one's self-interest, TEMPT is more likely to enter players' considerations when it is salient, that is, in Study 1 and less likely in Study 2.

RISK has no significant impact on cooperation in any of our three experiments. A likely reason is that for RISK to become relevant, people need to believe that their opponent is likely to defect in which case most people want to defect anyway.

7. Concluding remarks

The PD occupies a place of fundamental importance in social science research on cooperation as it represents the simplest setting in which individual and collective interests diverge. An extensive body of experimental research uses money payoffs to generate games where individuals maximize their own earnings by defecting, while combined earnings are

maximized by cooperating. This research shows that many individuals cooperate, even in oneshot games, but nevertheless the literature offers an incomplete account of how the money payoffs affect cooperation.

In this paper we examine the separate influences of the unilateral incentives to defect and the efficiency gains from cooperation. Following Mengel (2018) we use the index of RISK to measure the incentive to defect against a defector, the index of TEMPT to measure the incentive to defect against a cooperator, and the index of EFF to measure the efficiency gains from cooperation.

Taken together, the combined evidence from our studies suggests that in one-shot games (i) EFF robustly influences cooperation positively in all three studies; (ii) RISK does not influence cooperation systematically in any of the three studies; and (iii) TEMPT reduces cooperation in the within-subject designs. To probe the robustness of our results, we also looked at three related payoff indices: normalized loss (closely related to RISK); normalized gain (closely related to TEMPT), and the K-index (closely related to EFF). These related payoff indices share the same numerator with our respective index but have different denominators. The results are largely consistent with our main findings.

As in many previous studies, we find evidence of cooperation, even in carefully controlled anonymous one-shot games. But we emphasize that this cooperation is not random, it varies systematically with the material payoffs of the game, most robustly with EFF, that is, the difference in payoffs from mutual cooperation R and mutual defection P. This finding is consistent with several well-established behavioral motives and it implies that if a choice architect could influence the material payoffs of a prisoner's dilemma, they would promote mutually beneficial cooperation most effectively by increasing the benefits from mutual cooperation relative to the outcome of mutual defection rather than alleviating the risk of cooperation.

We conclude with a caveat. These results are obtained in the context of one-shot games. These are the building blocks of repeated games, but it remains an interesting question whether our results would carry over to a repeated game setting. In the repeated game there are likely to be learning effects as well as incentives for strategic cooperation that may impact our findings. We expect that initial cooperation would increase in EFF, and cooperation would be more difficult to sustain with higher TEMPT.

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Data availability

Data and analysis code are available at https://osf.io/mprsc/ [upon publication of the paper]

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Online Appendix to

The Role of Payoff Parameters for Cooperation in the One-Shot Prisoner's Dilemma

Simon Gächter^{1,2,3,†}, Kyeongtae Lee⁴, Martin Sefton¹ and Till O. Weber⁵

¹ Centre for Decision Research and Experimental Economics (CeDEx), University of Nottingham, UK

² IZA Bonn, Germany ³ CESifo Munich, Germany

⁴ Economic Research Institute, Bank of Korea, Korea

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⁵ Newcastle University Business School, Newcastle University, UK

[†]Corresponding author: simon.gaechter@nottingham.ac.uk

Appendix A. Instructions

Instructions for the preliminary experiment

You are now taking part in an economic experiment. Depending on the decisions made by you and other participants, you can earn a considerable amount of money. It is therefore very important that you read these instructions with care.

These instructions are solely for your private use. It is prohibited to communicate with other participants during the experiment. If you have any questions, please raise your hand. A member of the experiment team will come and answer them in private. If you violate this rule, you will be dismissed from the experiment and you will forfeit all payments.

You will solve several tasks during this experimental session. After this experimental session, one task will be randomly selected for payoff.

Additionally, you will receive a show-up fee of £3. Your earnings will be paid to you privately in cash at the end of the session.

At the end of the session, you will be asked to fill in a questionnaire. The answers you provide in this questionnaire are completely anonymous. They will not be revealed to anyone either during the experiment or after it. Furthermore, your responses to the questionnaires will not affect your earnings during the experiment.

You will be randomly matched with another participant. You will not learn who the other person, who you are matched with, at any point during or after the experiment.

The experiment

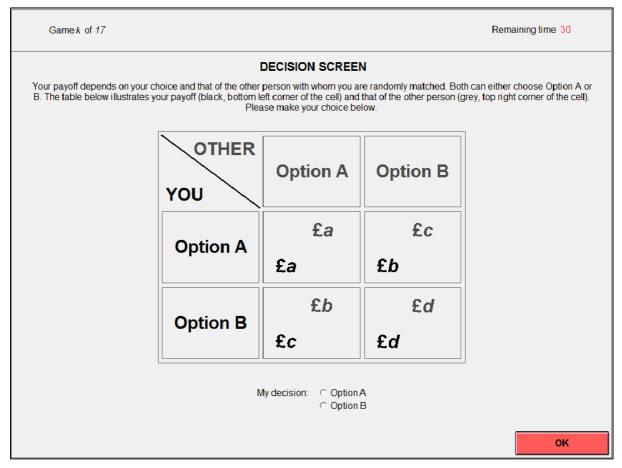
The experiment consists of 17 games and is separated into two stages: the decision stage and the results stage.

At the decision stage, you will have to make a decision for each of the 17 games. The other person, with whom you are randomly matched, will also make a decision for each of the 17 games. During the decision stage, you will not receive any feedback on the choices of the other person and the outcome of the games.

At the results stage, you will receive feedback on the decision taken by you, the other player's decision, as well as the resulting payoffs from these choices.

The decision stage

At the decision stage, you will see the following screen for each game:



In the table shown on the decision screen, your actions and resulting payoffs are given in black (bottom left corner) and the other person's actions and payoffs are given in grey (top right corner). The payoffs shown will be paid to you in case this game is randomly selected at the end of the session. The table is read as follows (black payoffs):

- If you choose Option A and the other participant chooses Option A, you receive £a.
- If you choose Option A and the other participant chooses Option B, you receive $\pounds b$.
- If you choose Option B and the other participant chooses Option A, you receive $\pounds c$.
- If you choose Option B and the other participant chooses Option B, you receive £d.

Note that the other participant (grey payoffs) is in the same situation as you are. The other participant will receive the following payoff, if this game is randomly selected at the end of the session:

- If the other participant chooses Option A and you choose Option A, the other participant receives £a.
- If the other participant chooses Option A you choose Option B, the other participant receives £b.

- If the other participant chooses Option B and you choose Option A, the other participant receives £c.
- If the other participant chooses Option B and you choose Option B, the other participant receives £d.

Keep in mind that you will not receive any feedback on the other person's choices and the other person's payoffs during the decision stage.

The results stage

The results stage starts after all participants have made their decisions for each of the 17 games. At the results stage you will learn the outcomes of each of the 17 games, starting with the first game. First, you will see the payoff table, with **your own choice** highlighted for several seconds. Afterwards, you will see the **other participant's choice and the resulting payoffs** for several seconds.

If you have any questions, please raise your hand and a member of the experiment team will come and answer them in private.

Instructions for Study 1

Note: These are the instructions used on Amazon Mechanical Turk. The instructions for the sessions conducted at the University of Nottingham used an exchange rate of 100 tokens = £1. Additionally, on the welcome screen, the term 'HIT' was replaced with 'experiment'. Otherwise, the instructions were identical.

Welcome

Thank you for accepting this HIT. To complete this HIT, you must make some decisions. Including the time for reading these instructions, the HIT will take about 30 minutes to complete. If you are using a desktop or laptop to complete this HIT, we recommend that you maximize your browser screen (press F11) before you start.

It is important that you complete this HIT without interruptions. During the HIT, please **do not close this window or get distracted from the task.** If you close your browser or leave the task, you will not be able to re-enter and we will not be able to pay you.

In this HIT, you will be matched with one other participant. Each of you will make decisions for 8 decision situations. In each situation, each of you will earn Tokens depending on your decisions.

At the end of the HIT, one of the decision situations will be randomly chosen. Your earnings from this situation will be converted from Tokens to Dollars at a rate of 100 Tokens = \$ 1. This will be added to **your participation fee of \$1.00**. Depending on your decisions, you may make up to \$8.00 more in addition to the \$1.00 participation fee. In the same way, Tokens earned by the person matched with you in that same situation will also be converted to Dollars at a rate of 100 Tokens = \$ 1.

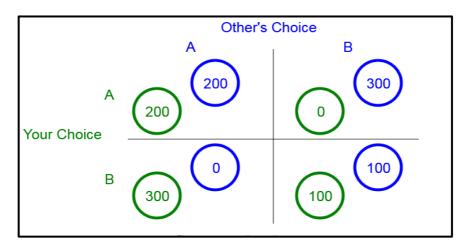
You will receive a code to collect your payment via MTurk upon completion.

Please click "Continue" to start the HIT.

Instructions

The HIT consists of 8 decision situations.

Each decision situation will be presented on a screen like the **example screen** below.



You and the other person will be making choices between **A** and **B**. Your earnings are the values in the green circle, and the other person's earnings are the values in the blue circle. The table is read as follows:

- If you choose A and the other person chooses A, you will earn 200 Tokens and the other person will earn 200 Tokens.
- If you choose A and the other person chooses B, you will earn 0 Tokens and the other person will earn 300 Tokens.
- If you choose B and the other person chooses A, you will earn 300 Tokens and the other person will earn 0 Tokens.
- If you choose B and the other person chooses B, you will earn 100 Tokens and the other person will earn 100 Tokens.

Please note that the values in the table will differ in each decision situation.

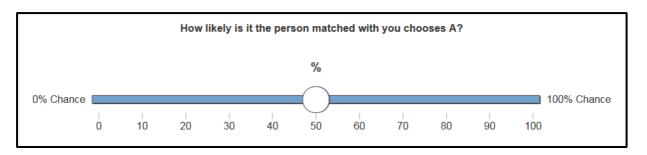
Tasks

In each decision situation, you must complete **two types** of tasks, which we will refer to below as the "decision" and "prediction".

• For the "decision" task, you will see the following screen and you must choose A or B:



• For the "prediction" task, you will see the following screen and you must indicate how likely you think it is that the other person will choose A:



During the HIT, you will not receive any feedback on the other person's choice or the outcomes of the decision situations.

Your dollar earnings

On completion of the HIT, you will be paid your participation fee of \$ 1.

In addition, one of the decision situations will be randomly chosen for your additional dollar earnings. Your earnings and the other person's earnings will be determined depending on choices of you and the other person in that situation. Two examples should make this clear.

Example 1. Assume that you choose A and the other person matched with you chooses A in the above example screen. As a consequence, you will earn 200 Tokens and the other person will earn 200 Tokens.

Example 2. Assume that you choose B and the other person matched with you chooses A in the above example screen. As a consequence, you will earn 300 Tokens and the other person will earn 0 Tokens.

At the end of the HIT

On completion of the HIT, one of the decision situations will be randomly chosen as explained above. You will be informed of your choices and earnings for that decision situation, and you will be paid these earnings in addition to your participation fee.

Note that we will not be able to pay you if you do not complete the HIT. If the person you are matched with does not complete the HIT, the computer will randomly select one of the four possible earnings in the randomly chosen decision situation, and you will be paid these earnings in addition to your participation fee.

Your participation fee and the additional earnings will be paid to you within two working days.

Instructions for Study 2

Welcome

Thank you for accepting this HIT. To complete this HIT, you must make some decisions. Including the time for reading these instructions, the HIT will take about 10 minutes to complete. If you are using a desktop or laptop to complete this HIT, we recommend that you maximize your browser screen (press F11) before you start.

It is important that you complete this HIT without interruptions. During the HIT, please **do not close this window or get distracted from the task.** If you close your browser or leave the task, you will not be able to re-enter and we will not be able to pay you.

In this HIT, you will be matched with one other participant. Each of you will make decisions for one decision situation. In this situation, each of you will earn Tokens depending on your decisions.

At the end of the HIT, your earnings from this situation will be converted from Tokens to Dollars at a rate of **250 Tokens** = \$ 1. This will be added to **your participation fee of \$1.50**. In the same way, Tokens earned by the person matched with you in that same situation will also be converted to Dollars at a rate of 250 Tokens = \$ 1.

You will receive a code to collect your payment via MTurk upon completion.

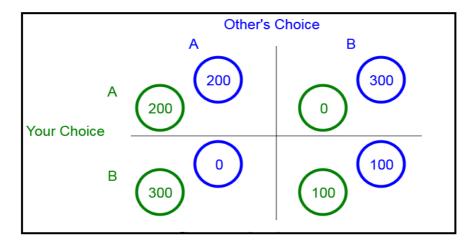
You will have to read some instructions and answer questions about them to make sure you understand the decision situation. You will have three attempts for the questions. If you fail to correctly answer the questions you will be removed from the HIT.

Please click "I DO want to Continue" to start the HIT.

Instructions

The HIT consists of one decision situation.

The decision situation will be presented on a screen like the **example screen** below.



You and the other person will be making choices between **A** and **B**. Your earnings are the values in the green circle, and the other person's earnings are the values in the blue circle. The table is read as follows:

- If you choose A and the other person chooses A, you will earn 200 Tokens and the other person will earn 200 Tokens.
- If you choose A and the other person chooses B, you will earn 0 Tokens and the other person will earn 300 Tokens.
- If you choose B and the other person chooses A, you will earn 300 Tokens and the other person will earn 0 Tokens.
- If you choose B and the other person chooses B, you will earn 100 Tokens and the other person will earn 100 Tokens.

Please note that the values in the table will differ in the actual decision situation.

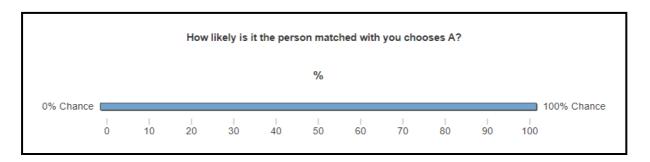
Tasks

In the decision situation, you must complete **two types** of tasks, which we will refer to below as the "decision" and "prediction".

• For the "decision" task, you will see the following screen and you must choose A or B:



• For the "prediction" task, you will see the following screen and you must indicate how likely you think it is that the other person will choose A:



Your dollar earnings

On completion of the HIT, you will be paid your participation fee of \$ 1.50 plus additional earnings based on your choice and the other person's choice in the decision situation. Two examples should make this clear.

Example 1. Assume that you choose A and the other person matched with you chooses A in the above example screen. As a consequence, you will earn 200 Tokens and the other person will earn 200 Tokens.

Example 2. Assume that you choose B and the other person matched with you chooses A in the above example screen. As a consequence, you will earn 300 Tokens and the other person will earn 0 Tokens.

At the end of the HIT

On completion of the HIT, you will be informed of your choices and earnings for the decision situation, and you will be paid these earnings in addition to your participation fee.

Note that we will not be able to pay you if you do not complete the HIT. If the person you are matched with does not complete the HIT, the computer will randomly select one of the four possible earnings in the decision situation, and you will be paid these earnings in addition to your participation fee.

Your participation fee and the additional earnings will be paid to you within two working days.

Appendix B. The preliminary experiment: details of design and results

This section provides the full details of the experimental design and procedures as well as all results of the preliminary experiment briefly described in Section 2.3 of the main text.

Experimental design and procedures

Our preliminary experiment used 17 PD games with payoff matrices similar to those reported in Mengel (2018) and Simpson (2003). Table B1 presents the payoff parameters for each of the 17 games (G1 to G17). G1 to G14 were designed by us, and G15 to G17 were taken from Simpson (2003). G1 to G15 meet the two standard prisoner's dilemma conditions (T > R > P > S and 2R > T + S). G16 and G17 do not satisfy the prisoner's dilemma conditions and are therefore excluded from the main analysis: G16 violates the first condition and G17 violates both conditions.¹

The two standard prisoner's dilemma conditions and non-negative payoff parameters restrict the theoretically possible variation of the payoff indices such that RISK \in (0, 1], TEMPT \in (0, 0.5) and EFF \in (0, 1). The implemented payoff parameters cover almost the entire possible range, with RISK varying from 0.04 to 1; TEMPT from 0.1 to 0.49; and EFF from 0.04 to 0.98. The design also includes several sets of games across which only one payoff index changes while holding the others constant. Games 1, 4, and 7 vary only in RISK. Games 2, 5, and 8 constitute a second set varying only in RISK. Games 10 and 11 vary only in TEMPT. Three sets of games vary only in EFF: Games 7, 10 and 13; Games 8, 11, 14; Games 9 and 12.

We ran our experiments with student participants at the University of Nottingham (n = 62). The subject pool characteristics are reported in Table C1 below. The experiment was computerized and conducted with z-Tree (Fischbacher (2007)). Participants were recruited using ORSEE (Greiner (2015)). None of the participants took part more than once. Participants were paired, made choices for all 17 games with no feedback between games and games presented in random order at the pair level. At the end of the session one game was chosen at random for each pair, and the pair were paid according to choices in this game. The sessions lasted for approximately one hour and the average earnings (including a £3 show-up fee) were £11.86 (SD = £3.32). Participants were paid in cash at the end of the session.

¹ We added the games by Simpson (2003) in the preliminary data collection because it is an early study with payoff variation (manipulating "fear" and "greed"). In Simpson (2003) the cooperation rates for G15 to G17 were, respectively, 0.45; 0.59; and 0.44 and they provide a benchmark for our results. These rates are similar to the ones we observed for G15 to G17 (of 0.56, 0.63, and 0.47, see Table 3). To keep the discussion focused, we did not

Table B1. Payoff parameters for the preliminary experiment.

| Game | T | R | P | S | RISK | TEMPT | EFF | Cooperation Rate |
|------|-----|------|-----|-----|------|-------|------|---------------------|
| G1 | 12 | 10.8 | 4.8 | 4.6 | 0.04 | 0.10 | 0.56 | 0.65 |
| G2 | 12 | 8.4 | 4.8 | 4.6 | 0.04 | 0.30 | 0.43 | 0.53 |
| G3 | 9.8 | 6.6 | 3.2 | 3 | 0.06 | 0.33 | 0.52 | 0.60 |
| G4 | 12 | 10.8 | 4.8 | 2.4 | 0.50 | 0.10 | 0.56 | 0.68 |
| G5 | 12 | 8.4 | 4.8 | 2.4 | 0.50 | 0.30 | 0.43 | 0.56 |
| G6 | 9.8 | 6.2 | 4.8 | 2.4 | 0.50 | 0.37 | 0.23 | 0.37 |
| G7 | 12 | 10.8 | 4.8 | 0 | 1.00 | 0.10 | 0.56 | 0.55 |
| G8 | 12 | 8.4 | 4.8 | 0 | 1.00 | 0.30 | 0.43 | 0.60 |
| G9 | 9.8 | 5 | 4.8 | 0 | 1.00 | 0.49 | 0.04 | 0.37 |
| G10 | 12 | 10.8 | 0.2 | 0 | 1.00 | 0.10 | 0.98 | 0.77 |
| G11 | 12 | 8.4 | 0.2 | 0 | 1.00 | 0.30 | 0.98 | 0.66 |
| G12 | 9.8 | 5 | 0.2 | 0 | 1.00 | 0.49 | 0.96 | 0.76 |
| G13 | 12 | 10.8 | 8 | 0 | 1.00 | 0.10 | 0.26 | 0.44 |
| G14 | 12 | 8.4 | 8 | 0 | 1.00 | 0.30 | 0.05 | 0.37 |
| G15 | 8 | 6 | 4 | 2 | 0.50 | 0.25 | 0.33 | 0.56 |
| G16 | 8 | 8 | 6 | 2 | 0.67 | 0 | 0.25 | 0.63 |
| G17 | 8 | 4 | 2 | 2 | 0 | 0.50 | 0.50 | 0.47 |

Notes: Payoffs in £. Games G1-G15 satisfy the two PD conditions T > R > P > S and 2R > T + S. Games G15 to G17 are taken from Simpson (2003) (with doubled payoffs relative to Simpson's original games). However, G16 and G17 do not satisfy the PD conditions: G16 violates the first condition and G17 violates both conditions. Cooperation Rate is the average cooperation rate we find in our experiment.

Results

Across the 15 PD games, cooperation rates vary from 0.37 to 0.77 (see Table B1). Of the participants, 81% were 'switchers' who altered their behavior at least once over the 15 games; 8% always defected; and 11% always cooperated. On average, participants cooperated in 8.47 of the 15 games. This suggests that the large variation in payoff indices implemented over the 15 games induced substantial variation in game play, and thus the impact of each payoff index warrants further investigation.

Fig. B1 plots the average cooperation rate in each of the 15 games against the respective RISK, TEMPT and EFF index. We find no significant association (at p < 0.10) between the average cooperation rate and RISK ($r_s = -0.00$, p = 0.992) or TEMPT ($r_s = -0.27$, p = 0.333).

However, the average cooperation rate is strongly positively and highly significantly correlated with EFF ($r_s = 0.90$, p < 0.001).

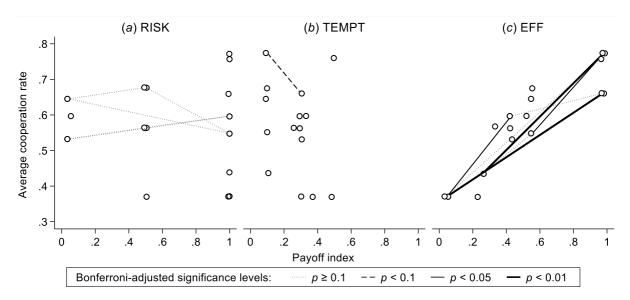


Fig. B1. Average cooperation rates and payoff indices of the 15 Prisoner's Dilemma games of the preliminary experiment. The line patterns indicate the Bonferroni-adjusted significance levels of a two-sided McNemar's test in pairs of games that vary only in a single payoff index.

In Fig. B1, pairs of games are connected by a line if one payoff index changes while the other two remain constant. The line pattern illustrates the Bonferroni-adjusted significance levels of a non-parametric McNemar's test for differences in the cooperation rates across a particular pair of games (see Table B2 for the uncorrected p-values). Panel (a) shows six pairs of games in which only RISK varies. We cannot reject the null hypothesis of equal cooperation rates across any of the pairs. Panel (b) shows one pair of games that varies only in TEMPT, and we find a weakly significantly lower cooperation rate associated with higher TEMPT. Panel (c) indicates seven pairs of games differing in EFF only, and we find substantial evidence of an effect of this index on behavior: we can (strongly) reject the null hypothesis of equal cooperation rates for five of the seven pair-wise comparisons possible.²

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² Table B1 in Appendix B reports the uncorrected p-values and confirms that our results are not driven by the correction for multiple testing.

Table B2. McNemar's tests for differences in cooperation across games in the preliminary experiment.

| | Games | Variation | Indices held constant | <i>p</i> -value |
|--------|------------|--------------|---------------------------|-----------------|
| RISK: | G1 vs G4 | 0.04 vs 0.50 | TEMPT = 0.10, EFF = 0.56 | 0.774 |
| | G2 vs G5 | 0.04 vs 0.50 | TEMPT = 0.30, EFF = 0.43 | 0.815 |
| | G4 vs G7 | 0.50 vs 1.00 | TEMPT = 0.10, EFF = 0.56 | 0.057 |
| | G5 vs G8 | 0.50 vs 1.00 | TEMPT = 0.30, EFF = 0.43 | 0.804 |
| | G1 vs G7 | 0.04 vs 1.00 | TEMPT = 0.10, EFF = 0.56 | 0.210 |
| | G2 vs G8 | 0.04 vs 1.00 | TEMPT = 0.30, EFF = 0.43 | 0.524 |
| TEMPT: | G10 vs G11 | 0.10 vs 0.30 | RISK = 1.00, EFF = 0.98 | 0.092* |
| EFF: | G7 vs G10 | 0.56 vs 0.98 | RISK = 1.00, TEMPT = 0.10 | 0.004** |
| | G10 vs G13 | 0.26 vs 0.98 | RISK = 1.00, TEMPT = 0.10 | < 0.001*** |
| | G7 vs G13 | 0.26 vs 0.56 | RISK = 1.00, TEMPT = 0.10 | 0.092 |
| | G8 vs G11 | 0.43 vs 0.98 | RISK = 1.00, TEMPT = 0.30 | 0.481 |
| | G11 vs G14 | 0.05 vs 0.98 | RISK = 1.00, TEMPT = 0.30 | < 0.001*** |
| | G8 vs G14 | 0.05 vs 0.43 | RISK = 1.00, TEMPT = 0.30 | 0.003** |
| | G9 vs G12 | 0.04 vs 0.96 | RISK = 1.00, TEMPT = 0.49 | < 0.001*** |

Notes: The column reports the uncorrected p-value from the McNemar's tests. To correct for multiple testing, we use Bonferroni-adjusted significance levels. The results from the Bonferroni correction are reported using the symbols in superscript. Bonferroni-adjusted significance levels for each payoff index are as follows. RISK: p < 0.017; ** p < 0.008; *** p < 0.002. TEMPT: *p < 0.1; ** p < 0.05; *** p < 0.01. EFF: *p < 0.014; **p < 0.007; *** p < 0.001.

In Table B3 Column 1, we report the full regression results corresponding to Table 3 of the main text. Column 2 shows that the results are robust to excluding games G9 and G14 (which involve extreme measures of EFF).³

13

³ We also ran regressions in which we included G16 and G17. These games violate the PD conditions and have multiple equilibria. If we include these games, that is, running the regression of Table 3 with all 17 games, we find that RISK remains insignificant [coeff (s.e.): -0.019 (0.034)]; TEMPT becomes significantly negative [-0.243 (0.082)] and EFF remains significantly positive [0.366 (0.060)].

Table B3. Payoff indices and cooperation in the preliminary experiment.

| Dependent variable: cooperation dummy | (1) 15 PD games | (2) Excluding G9 & G14 |
|---------------------------------------|---------------------|---------------------------|
| RISK | -0.044 (0.036) | -0.052 (0.045) |
| TEMPT | -0.083 (0.087) | -0.108 (0.090) |
| EFF | 0.399*** (0.060) | 0.415*** (0.077) |
| Round | -0.001 (0.003) | -0.001 (0.003) |
| Age | 0.002 (0.013) | 0.000 (0.013) |
| Female | 0.135 (0.103) | 0.120 (0.104) |
| Nationality: China | 0.055 (0.159) | 0.052 (0.162) |
| Nationality: Malaysia | 0.123 (0.119) | 0.125 (0.124) |
| Nationality: Other | -0.045 (0.136) | -0.039 (0.138) |
| Business/Economics major | -0.076 (0.092) | -0.091 (0.093) |
| Spending: Above median | 0.040 (0.106) | 0.054 (0.109) |
| Spending: Prefer not to say | -0.077 (0.109) | -0.099 (0.112) |
| Political attitude: Left | 0.078 (0.101) | 0.070 (0.104) |
| Political attitude: Right | 0.067 (0.165) | 0.040 (0.166) |
| Constant | 0.249 (0.340) | 0.303 (0.343) |
| Within R ² | 0.10 | 0.07 |
| Obs. (Clusters) | 930 (62) | 806 (62) |

Notes: G9 and G14 have extreme values of EFF (0.04 and 0.05, resp.). Coefficients of a random effects linear probability model with robust standard errors clustered on participants in parentheses. * p < 0.1; *** p < 0.05; **** p < 0.01.

Figure B2 illustrates the associations between the average cooperation rates in the prisoner's dilemma games (excluding G9 and G14, which have extreme values for l and g) and normalized loss l, normalized gain g. We find a highly significant negative association between cooperation and normalized loss ($r_s = -0.75$, p = 0.003), and a weakly significant negative association between cooperation and normalized gain ($r_s = -0.51$, p = 0.074).

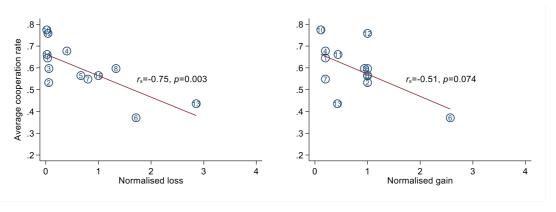


Fig. B2. Normalized loss l, normalized gain g, and average cooperation rates in the preliminary experiment. The game number is shown in the respective marker. G9 and G14 are excluded.

In Table B4, we estimate the effects of normalized payoff parameters on cooperation. Since l, g and the K-index are highly correlated, we run separate regression models to estimate their effects on cooperation. It is also important to note that, unlike in Studies 1 and 2, the prisoner's dilemma games included in the preliminary experiment do not provide an orthogonal variation in l and g.

Columns 1-4 show the effect of normalized loss l and normalized gain g on cooperation. For comparison, we exclude G9 and G14, which have extreme values for l and g, in Columns 2 and 4. We find highly significant effects of l and g on cooperation independent of the inclusion of games with extreme values for the normalized payoff parameters.

Figure B3 illustrates the association between average cooperation rates and the K-index in the preliminary experiment, which is highly significant and positive ($r_s = 0.83$, p < 0.001). Table B4, Column 5 show the results of a regression model to estimate the effect of the K-index on cooperation. We find highly significant positive effects of the K-index on cooperation.

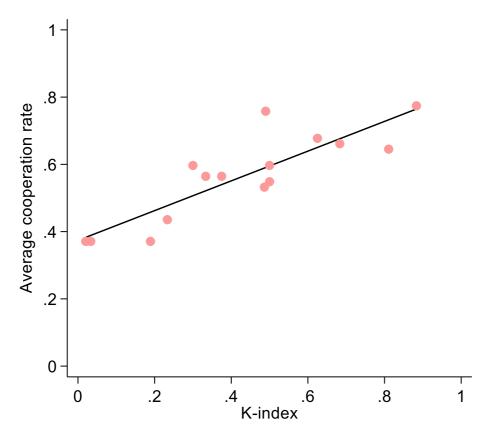


Fig. B3. Average cooperation rates for each level of a game's K-index in the preliminary experiment. *Note*: The black line shows the predicted values from a linear regression.

Table B4. Normalized loss l, normalized gain g, K-index and cooperation in the preliminary experiment.

| Dependent variable: | (1) | (2) | (3) | (4) | (5) |
|-----------------------------|----------------------|----------------------|----------------------|----------------------|---------------------|
| cooperation dummy | 15 PD games | Excluding G9 & G14 | 15 PD games | Excluding G9 & G14 | 15 PD games |
| Normalized loss l | -0.011*** (0.002) | -0.098*** (0.021) | | U14 | |
| Normalized gain g | | | -0.012*** (0.002) | -0.103*** (0.023) | |
| K-index | | | | | 0.442*** (0.065) |
| Round | -0.001 | 0.000 | -0.002 | -0.001 | -0.002 |
| | (0.002) | (0.003) | (0.002) | (0.003) | (0.003) |
| Age | 0.002 | 0.000 | 0.002 | 0.000 | 0.002 |
| | (0.013) | (0.013) | (0.013) | (0.013) | (0.013) |
| Female | 0.135 | 0.120 | 0.135 | 0.120 | 0.135 |
| | (0.103) | (0.104) | (0.103) | (0.104) | (0.103) |
| Nationality: China | 0.055 | 0.052 | 0.055 | 0.052 | 0.055 |
| | (0.159) | (0.162) | (0.159) | (0.162) | (0.159) |
| Nationality: | 0.123 | 0.125 | 0.123 | 0.125 | 0.123 |
| Malaysia | (0.119) | (0.124) | (0.119) | (0.124) | (0.119) |
| Nationality: Other | -0.045 | -0.039 | -0.045 | -0.039 | -0.045 |
| | (0.136) | (0.138) | (0.136) | (0.138) | (0.136) |
| Business/Eco- | -0.076 | -0.091 | -0.076 | -0.091 | -0.076 |
| nomics major | (0.092) | (0.093) | (0.092) | (0.093) | (0.092) |
| Spending: Above median | 0.040 | 0.054 | 0.040 | 0.054 | 0.040 |
| | (0.106) | (0.108) | (0.106) | (0.108) | (0.106) |
| Spending: Prefer not to say | -0.077 | -0.099 | -0.077 | -0.099 | -0.077 |
| | (0.109) | (0.112) | (0.109) | (0.112) | (0.109) |
| Political attitude: | 0.078 | 0.070 | 0.078 | 0.070 | 0.078 |
| Left | (0.101) | (0.104) | (0.101) | (0.104) | (0.101) |
| Political attitude: | 0.067 | 0.039 | 0.068 | 0.040 | 0.067 |
| Right | (0.165) | (0.166) | (0.165) | (0.166) | (0.165) |
| Constant | 0.428 | 0.535 | 0.429 | 0.558 | 0.202 |
| | (0.347) | (0.351) | (0.346) | (0.350) | (0.343) |
| Within R ² | 0.05 | 0.05 | 0.04 | 0.03 | 0.08 |
| Obs. (Clusters) | 930 (62) | 806 (62) | 930 (62) | 806 (62) | 930 (62) |

Notes: G9 and G14 have extreme values of g and l. Coefficients of a random effects linear probability model with robust standard errors clustered on participants in parentheses. * p < 0.1; *** p < 0.05; **** p < 0.01.

Appendix C. Descriptive statistics of subject pools

Table C1. Descriptive statistics of subject pools

| | Prelim. exp UoN | Study 1: UoN | Study 1: AMT | Study 2: AMT |
|--------------------------------------|--------------------|-----------------|-----------------|-----------------|
| Age (SD) | 22.76 (2.98) | 22.27(4.33) | 33.56(9.76) | 38.07(11.00) |
| Female | 74.19% | 60.38% | 49.59% | 45.14% |
| Nationality | | | | |
| United Kingdom | 16.13% | 61.63% | | |
| China | 19.35% | 5.66% | | |
| Malaysia | 40.32% | 7.55% | | |
| Others | 24.19% | 25.16% | | |
| Ethnicity | | | | |
| White | | 48.43% | 71.90% | 87.13% |
| Asian | | 27.04% | 11.57% | 2.98% |
| Black | | 6.92% | 7.44% | 5.78% |
| Others | | 17.61% | 7.44% | 3.60% |
| Highest level of education completed | | | | |
| Attended college | 69.36% | 72.33% | 21.49% | 12.31% |
| Undergraduate degree | 30.64% | 23.90% | 40.50% | 60.29% |
| Postgraduate degree | 0% | 0% | 20.66% | 20.20% |
| Business/Economics major | 25.81% | 20.75% | | |
| Income/spending* | | | | |
| Below median | 56.46% | 67.30% | 56.20% | 69.86% |
| Above median | 19.35% | 26.42% | 41.32% | 29.33% |
| Prefer not to say | 24.19% | 6.28% | 2.48% | 0.81% |
| Full-time worker | | | 58.12% | 82.85% |
| Political attitude | | | | |
| Left | 20.97% | 56.60% | 50.41% | 35.61% |
| Neither left nor right | 69.35% | 17.61% | 12.40% | 7.33% |
| Right | 9.68% | 25.79% | 37.19% | 57.05% |
| Experience in experiments (%) | | 49.06% | 38.01% | 41.21% |

Note: * In the preliminary experiment, participants were asked to state their monthly budget excluding accommodation. In Study 1, UoN participants were asked to choose one of the categories for reporting their average spending per month excluding a rent: Less than £200, £200-£399, £400-£599, £600-£799, £800-£999, £1,000 or more, prefer not to say. In Studies 1 and 2, AMT Participants were asked to choose one of the categories regarding their household pre-tax income: Less than \$30,000, \$30,000-\$49,999, \$50,000-\$69,999, \$70,000-\$89,999, \$90,000 or more, prefer not to say. Experience in experiments denotes the share of subjects who answered that they had previously participated in other experiments more than twice.

Comparing UoN samples: We compare subject pool differences across the preliminary experiment and Study 1's UoN sample. We find only small differences in the average age of participants, yet these differences are highly significant (Mann-Whitney Z = 3.24, p = 0.001). The share of females in the preliminary experiment is weakly significantly higher ($\chi^2(1) = 3.71$, p = 0.054). The distribution of nationalities is highly significantly different ($\chi^2(3) = 53.09$,

p < 0.001). The share of participants who study for a Business or Economics degree appears similar ($\chi^2(1) = 0.01$, p = 0.921). We do find a highly significant difference in participants' spending ($\chi^2(2) = 14.37$, p = 0.001). The political attitude also differs significantly ($\chi^2(2) = 54.77$, p < 0.001).

Comparing AMT samples: Next, we compare subject pool differences across Study 1's AMT sample and Study 2. We find highly significant differences in the participants' average age (Mann-Whitney Z = -4.66, p < 0.001). The share of females is comparable across samples $(\chi^2(1) = 0.93, \ p = 0.335)$. The distribution of ethnicities varies highly significantly $(\chi^2(3) = 31.28, \ p < 0.001)$. Education levels highly significantly differ across samples $(\chi^2(3) = 29.45, \ p < 0.001)$. We find a highly significant difference in participants' income $(\chi^2(2) = 11.81, p = 0.003)$. The political attitude differs significantly $(\chi^2(2) = 18.46, p < 0.001)$. The share of participants who have previously participated in experiment appears similar $(\chi^2(1) = 0.23, p = 0.633)$.

Appendix D. Non-parametric test results of Studies 1 and 2

Table D1. McNemar's tests for differences in cooperation across games in Study 1.

| | | UoN | AMT |
|-------------------|-----------------------|-----------------|-----------------|
| · | | <i>p</i> -value | <i>p</i> -value |
| Low vs High RISK | Low TEMPT, Low EFF | 0.188 | 0.860 |
| | Low TEMPT, High EFF | 0.489 | 0.851 |
| | High TEMPT, Low EFF | 1.000 | 0.743 |
| | High TEMPT, High EFF | 0.885 | 0.215 |
| Low vs High TEMPT | Low RISK, Low EFF | 0.014* | 0.441 |
| | Low RISK, High EFF | 0.012** | 0.034 |
| | High RISK, Low EFF | 0.296 | 0.296 |
| | High RISK, High EFF | 0.104 | < 0.001*** |
| Low vs High EFF | Low RISK, Low TEMPT | 0.016* | 0.061 |
| | Low RISK, High TEMPT | 0.092 | 0.775 |
| | High RISK, Low TEMPT | 0.006** | 0.024* |
| | High RISK, High TEMPT | 0.020* | 0.864 |

Notes: The columns report the uncorrected p-value from the McNemar's tests. To correct for multiple testing, we use Bonferroni-adjusted significance levels. The results from the Bonferroni correction are reported using the symbols in superscript. Bonferroni-adjusted significance levels are as follows. *p < 0.025; *** p < 0.013; **** p < 0.003.

Table D2. Fisher's exact tests for differences in cooperation across games in Study 2.

| | | <i>p</i> -value |
|-------------------|----------------------------------|-----------------|
| Low vs High RISK | Low TEMPT, Low EFF (G5 vs G6) | 0.549 |
| | Low TEMPT, High EFF (G1 vs G2) | 0.682 |
| | High TEMPT, Low EFF (G7 vs G8) | 0.616 |
| | High TEMPT, High EFF (G3 vs G4) | 0.228 |
| Low vs High TEMPT | Low RISK, Low EFF (G5 vs G7) | 0.072 |
| | Low RISK, High EFF (G1 vs G3) | 0.686 |
| | High RISK, Low EFF (G6 vs G8) | 0.618 |
| | High RISK, High EFF (G2 vs G4) | 0.042 |
| Low vs High EFF | Low RISK, Low TEMPT (G5 vs G1) | 0.004** |
| | Low RISK, High TEMPT (G7 vs G3) | 0.614 |
| | High RISK, Low TEMPT (G6 vs G2) | 0.009** |
| | High RISK, High TEMPT (G8 vs G4) | 1.000 |

Notes: The columns report the uncorrected p-value from Fisher's exact tests. To correct for multiple testing, we use Bonferroni-adjusted significance levels. The results from the Bonferroni correction are reported using the symbols in superscript. Bonferroni-adjusted significance levels are as follows. *p < 0.025; *** p < 0.013; **** p < 0.003.

Table D3. Wilcoxon signed-rank tests for differences in cooperative beliefs across games in Study 1.

| | | UoN <i>p</i> -value | AMT p-value |
|-------------------|-----------------------|------------------------|-------------|
| Low vs High RISK | Low TEMPT, Low EFF | 0.338 | 0.717 |
| | Low TEMPT, High EFF | 0.842 | 0.060 |
| | High TEMPT, Low EFF | 0.175 | 0.374 |
| | High TEMPT, High EFF | 0.658 | 0.775 |
| Low vs High TEMPT | Low RISK, Low EFF | 0.100 | 0.792 |
| | Low RISK, High EFF | 0.151 | 0.280 |
| | High RISK, Low EFF | 0.251 | 0.275 |
| | High RISK, High EFF | 0.016* | 0.041 |
| Low vs High EFF | Low RISK, Low TEMPT | 0.007** | 0.236 |
| | Low RISK, High TEMPT | < 0.001*** | 0.718 |
| | High RISK, Low TEMPT | 0.006** | < 0.001*** |
| | High RISK, High TEMPT | 0.034 | 0.066 |

Notes: The columns report the uncorrected p-value from the Wilcoxon signed-rank tests. To correct for multiple testing, we use Bonferroni-adjusted significance levels. The results from the Bonferroni correction are reported using the symbols in superscript. Bonferroni-adjusted significance levels are as follows. *p < 0.025; **p < 0.013; ***p < 0.003.

Table D4. Mann-Whitney tests for differences in cooperative beliefs across games in Study 2.

| | · | <i>p</i> -value |
|-------------------|----------------------------------|-----------------|
| Low vs High RISK | Low TEMPT, Low EFF (G5 vs G6) | 0.298 |
| | Low TEMPT, High EFF (G1 vs G2) | 0.249 |
| | High TEMPT, Low EFF (G7 vs G8) | 0.447 |
| | High TEMPT, High EFF (G3 vs G4) | 0.007** |
| Low vs High TEMPT | Low RISK, Low EFF (G5 vs G7) | 0.087 |
| | Low RISK, High EFF (G1 vs G3) | 0.995 |
| | High RISK, Low EFF (G6 vs G8) | 0.907 |
| | High RISK, High EFF (G2 vs G4) | 0.141 |
| Low vs High EFF | Low RISK, Low TEMPT (G5 vs G1) | 0.011** |
| | Low RISK, High TEMPT (G7 vs G3) | 0.313 |
| | High RISK, Low TEMPT (G6 vs G2) | 0.686 |
| | High RISK, High TEMPT (G8 vs G4) | 0.305 |

Notes: The columns report the uncorrected p-value from Mann-Whitney tests. To correct for multiple testing, we use Bonferroni-adjusted significance levels. The results from the Bonferroni correction are reported using the symbols in superscript. Bonferroni-adjusted significance levels are as follows. *p < 0.025; **p < 0.013; *** p < 0.003.

Appendix E. Belief accuracy in Studies 1 and 2

Study 1. We measure belief accuracy as the absolute difference between individual beliefs and actual average cooperation rates in each game (UoN: 0.250, SD = 0.106, AMT: 0.268, SD = 0.099; individuals as independent observation). Thus, our measure of belief accuracy can be interpreted as an individual's belief deviation from the actual average cooperation rate in percentage points. On average, we find that UoN subjects reveal slightly higher belief accuracy, but the difference is only marginally significant (Mann-Whitney Z = -1.84, p = 0.066; individuals as independent observation). Table E1 documents belief accuracy by game and subject pool.

Table E1. Belief accuracy across games and samples in Study 1.

| | G1 | G2 | G3 | G4 | G5 | G6 | G7 | G8 |
|------------------------------|---------|---------|---------|---------|---------|---------|---------|---------|
| UoN | 0.255 | 0.253 | 0.251 | 0.243 | 0.263 | 0.259 | 0.221 | 0.255 |
| | (0.144) | (0.157) | (0.174) | (0.154) | (0.154) | (0.181) | (0.174) | (0.195) |
| AMT | 0.274 | 0.263 | 0.251 | 0.279 | 0.278 | 0.257 | 0.268 | 0.273 |
| | (0.151) | (0.144) | (0.145) | (0.164) | (0.150) | (0.137) | (0.154) | (0.151) |
| Mann-Whitney <i>p</i> -value | 0.316 | 0.550 | 0.565 | 0.078 | 0.274 | 0.382 | 0.003 | 0.074 |

Note. Belief accuracy is defined as the absolute deviation of beliefs from cooperation rates. *SD* in parentheses.

Study 2. Belief accuracy is defined as described above. Table E2 documents belief accuracy by game.

Table E2. Belief accuracy across games in Study 2.

| G1 | G2 | G3 | G4 | G5 | G6 | G7 | G8 |
|------------------|------------------|------------------|------------------|------------------|---------------|------------------|---------------|
| 0.317 (0.161) | 0.325 (0.178) | 0.300 (0.144) | 0.317 (0.151) | 0.308 (0.148) | 0.314 (0.136) | 0.309 (0.140) | 0.319 (0.139) |

Note. Belief accuracy is defined as the absolute deviation of beliefs from cooperation rates. *SD* in parentheses.

Appendix F. Full regression results of Studies 1 and 2

Table F1. Payoff indices, beliefs, and cooperation in Study 1.

| Ia | Table F1. Payoff indices, beliefs, and cooperation in Study 1. | | | | | |
|--------------------------|---|-----------------|-----------------|----------------|-----------------|-----------|
| Dependent | (1) | (2) | (3) | (4) | (5) | (6) |
| variable: | UoN | AMT | UoN | AMT | UoN | AMT |
| variable. | Coop. | Coop. | Belief | Belief | Coop. | Coop. |
| RISK | -0.094 | -0.073 | 0.001 | -0.014 | -0.094 | -0.066 |
| KISK | (0.062) | (0.062) | (0.038) | (0.051) | (0.058) | (0.062) |
| TEMPT | -0.408*** | -0.506*** | -0.147** | -0.174** | -0.323*** | -0.431*** |
| TEMPT | (0.108) | (0.117) | (0.062) | (0.078) | (0.107) | (0.112) |
| DDD | 0.245*** | 0.145** | 0.155*** | 0.076* | 0.155*** | 0.113 |
| EFF | (0.058) | (0.073) | (0.033) | (0.042) | (0.059) | (0.072) |
| D 11 0 | | | | | 0.582*** | 0.433*** |
| Belief | | | | | (0.061) | (0.068) |
| | -0.023*** | -0.014** | -0.015*** | -0.012*** | -0.014*** | -0.009 |
| Round | (0.006) | (0.006) | (0.003) | (0.004) | (0.005) | (0.006) |
| | 0.005 | -0.024 | -0.065*** | -0.157*** | 0.044* | 0.048 |
| B_is_Coop | (0.024) | (0.033) | (0.019) | (0.025) | (0.025) | (0.031) |
| | 0.034 | -0.026 | -0.003 | 0.004 | 0.036* | -0.028 |
| BeliefThenChoice | (0.023) | (0.027) | (0.014) | (0.015) | (0.022) | (0.026) |
| Age | 0.008 | 0.006* | 0.001 | 0.005** | 0.008 | 0.004 |
| | (0.007) | (0.003) | (0.003) | (0.002) | (0.006) | (0.003) |
| Female | 0.014 | 0.094 | 0.006 | 0.029 | 0.011 | 0.081 |
| | (0.057) | (0.065) | (0.034) | (0.023) | (0.049) | (0.056) |
| | 0.037 | 0.093 | -0.024 | 0.077 | 0.051 | 0.059 |
| Ethnicity: Asian | (0.064) | (0.099) | (0.038) | (0.068) | (0.054) | (0.085) |
| | 0.072 | 0.042 | 0.081* | 0.056 | 0.025 | 0.017 |
| Ethnicity: Black | (0.086) | (0.101) | (0.042) | (0.052) | (0.082) | (0.091) |
| | , , | 0.014 | ` ′ | , , | ` , | , , |
| Ethnicity: Other | 0.002 | (0.131) | 0.021 | 0.003 | -0.010 | 0.013 |
| , | (0.083) | , | (0.045) | (0.077) | (0.067) | (0.111) |
| Business/Economics | -0.101* | | -0.082** | | -0.054 | |
| major | (0.061) | | (0.033) | | (0.053) | |
| Spending/Income: | 0.072 | -0.079 | 0.080** | -0.031 | 0.025 | -0.066 |
| Above median | (0.060) | (0.065) | (0.039) | (0.036) | (0.049) | (0.055) |
| Spending/Income: | 0.136 | -0.164 | 0.043 | -0.064 | 0.111 | -0.137 |
| Prefer not to say | (0.122) | (0.214) | (0.073) | (0.156) | (0.097) | (0.151) |
| Political attitude: | 0.234*** | -0.127 | 0.055 | -0.078 | 0.202*** | -0.094 |
| Left | (0.065) | (0.095) | (0.042) | (0.061) | (0.056) | (0.078) |
| Political attitude: | 0.180** | -0.103 | 0.047 | -0.085 | 0.152** | -0.066 |
| Right | (0.077) | (0.091) | (0.046) | (0.059) | (0.068) | (0.075) |
| Previous experience | -0.067 | 0.012 | -0.022 | -0.011 | -0.054 | 0.017 |
| in experiments | (0.052) | (0.069) | (0.030) | (0.037) | (0.042) | (0.058) |
| | 0.180 | 0.575*** | 0.476*** | 0.552*** | -0.098 | 0.334** |
| Constant | (0.193) | (0.155) | (0.098) | (0.105) | (0.164) | (0.139) |
| Within R^2 | 0.06 | 0.04 | 0.07 | 0.13 | 0.15 | 0.06 |
| Obs. (Clusters) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) |
| Notes: Coefficients of a | random effec | ts linear proba | bility model (C | ols 1-2 5-6) c | or linear model | |

Notes: Coefficients of a random effects linear probability model (Cols. 1-2, 5-6) or linear model (Col. 3-4) with robust standard errors clustered on participants in parentheses. * p < 0.1; ** p < 0.05; *** p < 0.01.

Table F2. Payoff indices, beliefs, and cooperation in Study 2.

| Dependent variable: | (1) | (2) | (3) |
|------------------------------------|-------------|-----------|---------------------|
| | Cooperation | Belief | Cooperation |
| RISK | -0.012 | -0.042 | 0.012 |
| | (0.067) | (0.030) | (0.065) |
| TEMPT | -0.032 | -0.029 | -0.015 |
| | (0.113) | (0.050) | (0.110) |
| EFF | 0.179*** | 0.048* | 0.152*** |
| | (0.059) | (0.026) | (0.057) |
| Belief | | | 0.565*** (0.053) |
| B_is_Coop | -0.309*** | -0.550*** | 0.002 |
| | (0.024) | (0.011) | (0.038) |
| BeliefThenChoice | -0.034 | 0.014 | -0.042* |
| | (0.024) | (0.011) | (0.023) |
| Age | 0.001 | 0.000 | 0.001 |
| | (0.001) | (0.000) | (0.001) |
| Female | 0.001 | 0.004 | -0.001 |
| | (0.024) | (0.011) | (0.024) |
| Ethnicity: Asian | -0.073 | 0.067 | -0.111* |
| | (0.073) | (0.045) | (0.060) |
| Ethnicity: Black | 0.086* | 0.036 | 0.066 |
| | (0.048) | (0.025) | (0.046) |
| Ethnicity: Other | 0.038 | 0.049 | 0.010 |
| | (0.065) | (0.030) | (0.061) |
| Spending/Income: | -0.021 | -0.016 | -0.012 |
| Above median | (0.026) | (0.012) | (0.026) |
| Spending/Income: | 0.036 | 0.074 | -0.006 |
| Prefer not to say | (0.157) | (0.063) | (0.146) |
| Political attitude: | 0.064 | 0.020 | 0.053 |
| Left | (0.052) | (0.030) | (0.046) |
| Political attitude: | 0.032 | 0.030 | 0.016 |
| Right | (0.051) | (0.029) | (0.046) |
| Previous experience in experiments | 0.036 | -0.017 | 0.045* |
| | (0.025) | (0.011) | (0.024) |
| Constant | 0.588*** | 0.771*** | 0.153 |
| | (0.094) | (0.047) | (0.097) |
| R^2 | 0.11 | 0.64 | 0.17 |
| Obs. | 1601 | 1601 | 1601 |

Notes: Coefficients of a linear probability model (Cols. 1, 3) or linear model (Col. 2) with robust standard errors in parentheses. * p < 0.1; *** p < 0.05; **** p < 0.01.

Table F3. Regression of cooperation on payoff indices, with payoff scale interaction effects in Study 1.

| | <u> 1n</u> | i Study 1. | | |
|---------------------------------|------------|------------|------------|-----------|
| Dependent variable: | (1) | (2) | (3) | (4) |
| cooperation dummy | UoN | AMT | UoN | AMT |
| RISK | -0.120 | -0.051 | -0.140* | -0.022 |
| KISK | (0.080) | (0.089) | (0.075) | (0.088) |
| TEMDT | -0.378*** | -0.258 | -0.253* | -0.215 |
| TEMPT | (0.133) | (0.164) | (0.130) | (0.157) |
| EEE | 0.190 | 0.558** | 0.079 | 0.568** |
| EFF | (0.253) | (0.268) | (0.237) | (0.271) |
| DIGIZALI' 1 EEE 1 | 0.053 | -0.043 | 0.093 | -0.089 |
| RISK*High EFF dummy | (0.116) | (0.133) | (0.106) | (0.135) |
| | -0.060 | -0.496** | -0.139 | -0.433* |
| TEMPT*High EFF dummy | (0.184) | (0.252) | (0.176) | (0.240) |
| | (0.10.) | (0.202) | 0.584*** | 0.431*** |
| Belief | | | (0.061) | (0.068) |
| | -0.023*** | -0.014** | -0.014*** | -0.009 |
| Round | (0.006) | (0.006) | (0.005) | (0.006) |
| | 0.006 | -0.024 | 0.045* | 0.048 |
| B_is_Coop | (0.024) | (0.034) | (0.025) | (0.031) |
| | 0.033 | -0.026 | 0.036* | -0.028 |
| BeliefThenChoice | (0.023) | (0.027) | (0.022) | (0.026) |
| | 0.008 | 0.006* | 0.008 | 0.004 |
| Age | (0.007) | | (0.006) | |
| | , | (0.003) | , | (0.003) |
| Female | 0.014 | 0.094 | 0.011 | 0.081 |
| | (0.057) | (0.065) | (0.049) | (0.056) |
| Ethnicity: Asian | 0.037 | 0.093 | 0.051 | 0.059 |
| Ž | (0.064) | (0.100) | (0.054) | (0.085) |
| Ethnicity: Black | 0.072 | 0.042 | 0.024 | 0.017 |
| , | (0.086) | (0.101) | (0.083) | (0.091) |
| Ethnicity: Other | 0.002 | 0.014 | -0.010 | 0.013 |
| 24444 | (0.083) | (0.131) | (0.067) | (0.111) |
| Business/Economics major | -0.101* | | -0.054 | |
| · · | (0.061) | | (0.053) | |
| Spending/Income: Above | 0.072 | -0.079 | 0.025 | -0.066 |
| median | (0.060) | (0.065) | (0.049) | (0.055) |
| Spending/Income: Prefer not | 0.136 | -0.164 | 0.111 | -0.137 |
| to say | (0.122) | (0.214) | (0.097) | (0.152) |
| Political attitude: Left | 0.234*** | -0.127 | 0.202*** | -0.094 |
| ronnear annude. Len | (0.065) | (0.095) | (0.056) | (0.078) |
| Dolitical attitudas Diaht | 0.180** | -0.103 | 0.152** | -0.066 |
| Political attitude: Right | (0.077) | (0.092) | (0.068) | (0.075) |
| Previous experience in | -0.067 | 0.012 | -0.054 | 0.017 |
| experiments | (0.052) | (0.069) | (0.042) | (0.058) |
| • | 0.201 | 0.409** | -0.071 | 0.153 |
| Constant | (0.207) | (0.181) | (0.186) | (0.169) |
| Within R ² | 0.06 | 0.04 | 0.15 | 0.07 |
| Obs. (Clusters) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) |
| Notes: Coefficients of a randon | \ / | | · / | · / |

Notes: Coefficients of a random effects linear probability model with robust standard errors clustered on participants in parentheses. * p < 0.1; *** p < 0.05; **** p < 0.01.

We also ran logit regressions for all studies and obtained very similar results. The data and do-files that include these commands are available at https://osf.io/mprsc/.

Appendix G. Individual beliefs and cooperative choice in Study 1

Fig. G1 shows that for both, the UoN and AMT samples, the average beliefs of participants increase with their number of cooperative choices across the eight games of Study 1. Investigating the participants' belief-choice combinations in Study 1 can help infer their underlying preferences. A positive correlation between beliefs and choices is consistent with preferences for conditional cooperation, that is, an increased inclination to cooperate the higher the expected cooperation of others is.

However, the box plots show substantial heterogeneity in beliefs for participants who never or always cooperate. This implies that the behavior of some participants may be explained by motivations other than conditional cooperation. For example, participants who never cooperate may do so regardless of the belief they hold about the likelihood of their opponent cooperating or because they think their opponent will defect. Some participants who cooperate in all eight games do not expect the same behavior in others, that is, they have relatively low average beliefs. Therefore, the behavior of these participants might be best explained by altruism or Kantian preferences rather than conditional cooperation.

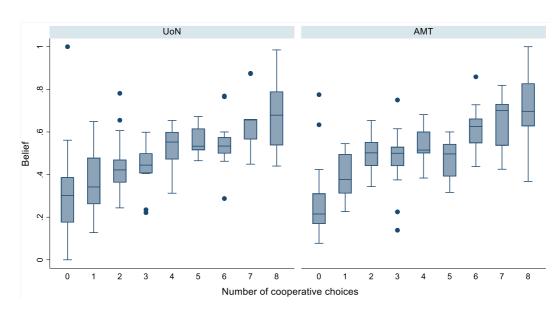


Fig. G1. The distribution of individuals' average beliefs by the number of their own cooperative choices (0 = defect in all 8 games; 8 = cooperate in all 8 games) for the UoN and AMT samples.

In addition to the aggregate level, we can also test conditional cooperation on the individual level by focusing on participants who cooperated in four out of the eight games (n = 32; Pooling the UoN and AMT samples). Conditional cooperation suggests that people's beliefs should be significantly higher in games for which they chose to cooperate compared to

games in which they defected. Our data provides evidence for this reasoning. For participants who cooperated exactly four times, we find significantly higher beliefs in the games for which they cooperated compared to the games in which they defected ($M_{\text{Defect}} = 0.46$, $M_{\text{Cooperate}} = 0.61$; Mann-Whitney Z = -3.99, p < 0.001).

Appendix H. The mediating role of beliefs for cooperation in Studies 1 and 2

Figure G1 shows the regression results as a path diagram. The path coefficients are taken from Table F1-F2.

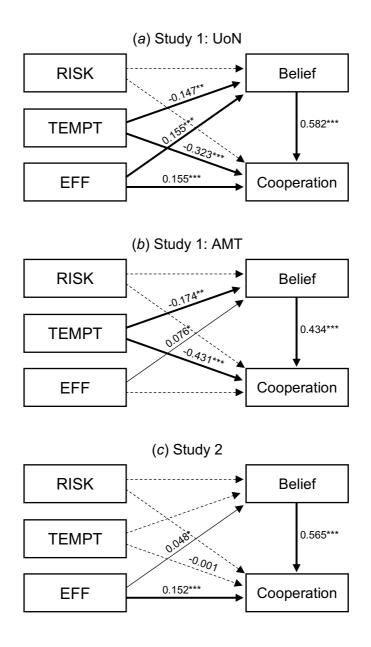


Fig. H1. Unstandardized path coefficient estimates taken from Table F1-F2. *Notes*: Control variables included in the regressions. Statistically non-significant coefficients are excluded. * p < 0.1; ** p < 0.05; *** p < 0.01.

Our mediation analysis and effect decomposition follow Baron and Kenny (1986). In a first step, we check if the effect of payoff indices on cooperation is plausibly mediated through beliefs. We find that beliefs are indeed a plausible mediator because (i) beliefs and cooperation are highly correlated (Study 1 UoN: $r_s = 0.48$, p < 0.001; Study 1 AMT: $r_s = 0.41$, $p_s < 0.001$; Study 2: $r_s = 0.38$, $p_s < 0.001$; Pooled samples), and (ii) when payoff indices significantly affect cooperation, they also significantly affect beliefs. For example, Table F1, Column 1 shows a highly significant total effect of TEMPT on cooperation in the UoN sample of Study 1 (c = 0.408, p < 0.001) and similarly Column 3 shows a significant direct effect of TEMPT on belief (a = -0.147, p = 0.017). Additionally, in the model reported in Column 5 that includes Belief as an explanatory variable, we see a reduction in the size of the coefficient of TEMPT.

Following Baron and Kenny (1986), the total effect of TEMPT on cooperation in the UoN sample of Study 1 (labelled c above), can therefore be decomposed into an indirect effect mediated through belief and a direct effect (c' = -0.323; Table F1, Col. 5). The indirect effect ab = -0.086 can be calculated as the product of the direct effect of TEMPT on the mediator belief (a = -0.147; Table F1, Col. 3) and the direct effect of the mediator belief on cooperation (b = 0.582; Table F1, Col. 5).

The standard error of the indirect effect can be approximated using a variant of the delta method, with

$$se_{ab} = \sqrt{b^2 se_a^2 + a^2 se_b^2 + se_a^2 se_b^2}.$$

In the current example the standard error of the indirect effect is 0.037. Now the z-value can be imputed as follows:

$$z = \frac{ab}{\sqrt{se_{ab}}}.$$

Assuming a standard normal distribution, the p-value can be calculated. In the current example, p = 0.011. The indirect effect of TEMPT on cooperation mediated through belief is thus negative and significant. It accounts for 21% of the total effect (0.086/0.408 = 0.211).

Appendix I. Monotonicity violations in Study 1

Figure I1 shows the frequency of violations of the monotonicity assumption separately for each payoff index in Study 1. This suggests that very few subjects violate the monotonicity assumption more than once for any given index.

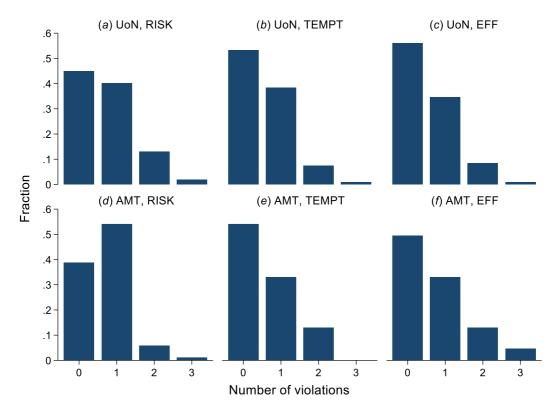


Fig. I1. The number of violations of the monotonicity assumption by payoff index in Study 1.

Note: To count the number of monotonicity violations, we compare twelve pairs of games: (*i*) 4 pairs of games where RISK changes given other payoff indices constant (i.e., G1 vs G2, G3 vs G4, G5 vs G6, G7 vs G8); (*ii*) 4 pairs of games where TEMPT changes given other payoff indices constant (i.e., G1 vs G3, G2 vs G4, G5 vs G7, G6 vs G8); (*iii*) 4 pairs of games of where EFF changes given other payoff indices constant (i.e., G1 vs G5, G2 vs G6, G3 vs G7, G4 vs G8). For an explanation of how violations are calculated see the main text, Section 4.2.

When aggregating the number of violations of the monotonicity assumptions across the three payoff indices for each individual, we find that 24% in UoN (16% in AMT) never violate the monotonicity assumption (Fig. I2). 17% in UoN (20% in AMT) display one violation and 59% in UoN (64% in AMT) violate the monotonicity assumption more than once when aggregating over the three indices.

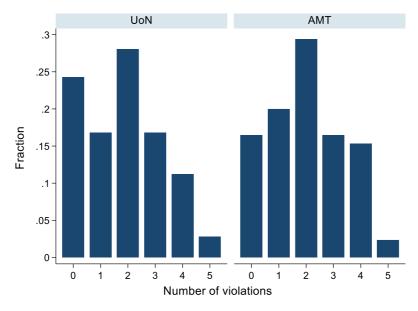


Fig. 12. The number of violations of the monotonicity assumption across all three payoff indices in Study 1.

Appendix J. Normalized loss, normalized gain, and the K-index in Studies 1 and 2

Table J1 summarizes the values of normalized loss l, normalized gain g and the K-index as implied by the game parameters we used in G1 to G8 in Studies 1 and 2. For further details see Section 5 in the main text. The results on l and g are in the main text Section 5.1 and the results on the K-index are in Section 5.2.

Table J1. Normalized loss, normalized gain, and K-index

| Game | T | R | P | S | Normalized loss <i>l</i> | Normalized gain g | K-index |
|------|-----|-----|-----|-----|---------------------------|---------------------------|---------------------------|
| | | | | | $l = \frac{P - S}{R - P}$ | $g = \frac{T - R}{R - P}$ | $K = \frac{R - P}{T - S}$ |
| G1 | 600 | 500 | 200 | 90 | 0.37 | 0.33 | 0.59 |
| G2 | 600 | 500 | 200 | 20 | 0.60 | 0.33 | 0.52 |
| G3 | 800 | 500 | 200 | 90 | 0.37 | 1.00 | 0.42 |
| G4 | 800 | 500 | 200 | 20 | 0.60 | 1.00 | 0.38 |
| G5 | 600 | 500 | 400 | 180 | 2.20 | 1.00 | 0.24 |
| G6 | 600 | 500 | 400 | 40 | 3.60 | 1.00 | 0.18 |
| G7 | 800 | 500 | 400 | 180 | 2.20 | 3.00 | 0.16 |
| G8 | 800 | 500 | 400 | 40 | 3.60 | 3.00 | 0.13 |

Figure J1 illustrates the variation in payoff parameters across the 8 games. Panel (a) shows the orthogonal variation in RISK, TEMPT and EFF. Panel (b) illustrates the variation in normalized loss l and normalized gain g for the four low-EFF games and the four high-EFF games.

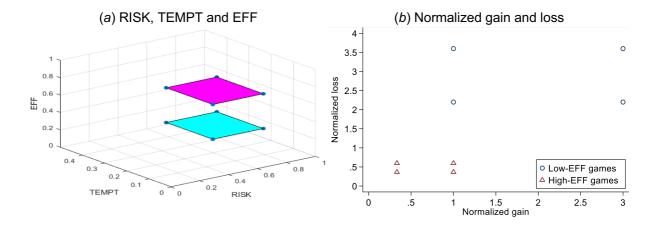


Fig. J1. The variation in RISK, TEMPT and EFF (Panel a) and normalized loss l and gain g (Panel b) across the eight games.

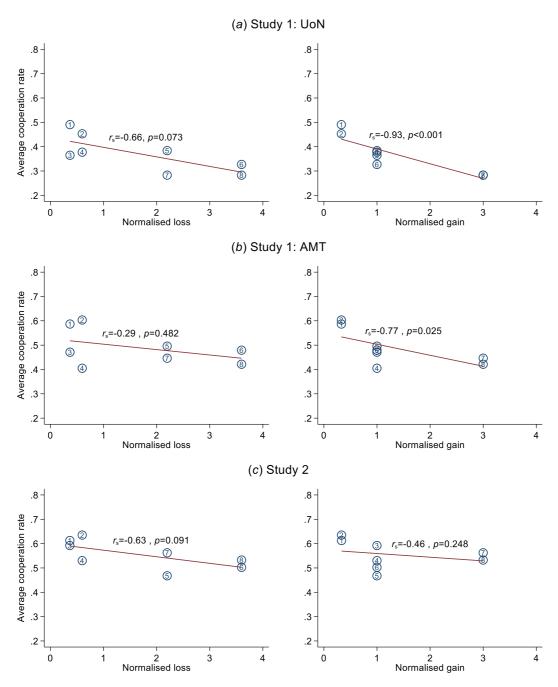


Fig. J2. Normalized loss l, normalized gain g, and average cooperation rates by study and subject pool. The game number is shown in the respective marker.

Table J2. Normalized loss l, normalized gain g, beliefs, and cooperation in the low-EFF games of Study 1.

| | | | Study 1. | | | |
|---------------------------------------|---------------------------------------|------------------------------|----------------------------------|----------------------------------|---------------------------------------|---------------------------------------|
| Dependent variable: cooperation dummy | (1) Study 1: UoN Cooperation | (2) Study 1: AMT Cooperation | (3) Study 1: UoN Belief | (4) Study 1: AMT Belief | (5) Study 1: UoN Cooperation | (6) Study 1: AMT Cooperation |
| | | | | | | |
| Normalized loss l | -0.030 | -0.011 | 0.008 | -0.017 | -0.034* | -0.002 |
| | (0.020) | (0.022) | (0.013) | (0.018) | (0.019) | (0.023) |
| Normalized gain g | -0.039*** | -0.025 | -0.023** | -0.010 | -0.026* | -0.020 |
| | (0.014) | (0.017) | (0.009) | (0.012) | (0.014) | (0.016) |
| Belief | | | | | 0.571*** (0.082) | 0.529*** (0.083) |
| Round | -0.022*** | -0.002 | -0.017*** | -0.004 | -0.012* | 0.002 |
| | (0.007) | (0.008) | (0.004) | (0.005) | (0.007) | (0.008) |
| B_is_Coop | 0.006 | 0.017 | -0.078*** | -0.146*** | 0.051 | 0.095** |
| | (0.030) | (0.042) | (0.022) | (0.032) | (0.031) | (0.040) |
| BeliefThenChoice | 0.022 | -0.031 | -0.002 | 0.024 | 0.025 | -0.044 |
| | (0.031) | (0.040) | (0.019) | (0.023) | (0.030) | (0.036) |
| Age | 0.010 | 0.008** | 0.002 | 0.006*** | 0.009 | 0.005 |
| | (0.008) | (0.004) | (0.004) | (0.002) | (0.007) | (0.003) |
| Female | 0.036 | 0.113 | 0.011 | 0.007 | 0.030 | 0.109* |
| | (0.064) | (0.072) | (0.038) | (0.036) | (0.057) | (0.061) |
| Ethnicity: Asian | 0.070 | 0.089 | -0.015 | 0.048 | 0.078 | 0.064 |
| | (0.069) | (0.106) | (0.041) | (0.068) | (0.061) | (0.086) |
| Ethnicity: Black | 0.143 | -0.025 | 0.127** | 0.077 | 0.070 | -0.067 |
| | (0.102) | (0.116) | (0.058) | (0.068) | (0.107) | (0.113) |
| Ethnicity: Other | -0.010 | -0.026 | -0.001 | 0.006 | -0.010 | -0.030 |
| | (0.098) | (0.130) | (0.052) | (0.087) | (0.078) | (0.103) |
| Business/Eco- nomics major | -0.123* (0.066) | | -0.107*** (0.035) | | -0.062 (0.061) | |
| Spending/Income: | 0.072 | -0.069 | 0.098** | -0.031 | 0.016 | -0.053 |
| Above median | (0.071) | (0.067) | (0.044) | (0.039) | (0.059) | (0.057) |
| Spending/Income: | 0.122 | -0.128 | 0.078 | -0.057 | 0.078 | -0.098 |
| Prefer not to say | (0.146) | (0.214) | (0.081) | (0.076) | (0.127) | (0.178) |
| Political attitude: | 0.226*** | -0.140 | 0.063 | -0.055 | 0.189*** | -0.111 |
| Left | (0.067) | (0.104) | (0.045) | (0.070) | (0.061) | (0.084) |
| Political attitude: | 0.172** | -0.090 | 0.060 | -0.090 | 0.138* | -0.043 |
| Right | (0.082) | (0.100) | (0.052) | (0.068) | (0.076) | (0.080) |
| Previous experience in experiments | -0.076 | 0.007 | -0.011 | -0.045 | -0.070 | 0.031 |
| | (0.057) | (0.076) | (0.033) | (0.043) | (0.049) | (0.065) |
| Constant | 0.178 | 0.345** | 0.464*** | 0.499*** | -0.090 | 0.074 |
| | (0.223) | (0.173) | (0.112) | (0.121) | (0.190) | (0.146) |
| Within R^2 | 0.05 | 0.01 | 0.08 | 0.10 | 0.12 | 0.04 |
| Obs. (Clusters) | 616 (154) | 476 (119) | 616 (154) | 476 (119) | 616 (154) | 476 (119) |
| Notes: Coefficients of a rar | \ / | | . , | \ / | \ / | |

Notes: Coefficients of a random effects linear probability model (Cols. 1-2, 5-6) or linear model (Col. 3-4) with robust standard errors clustered on participants in parentheses. * p < 0.1; ** p < 0.05; *** p < 0.01.

Table J3. Normalized loss l, normalized gain g, beliefs, and cooperation in the low-EFF games of Study 2.

| Dependent variable: | (1) | (2) | (3) |
|------------------------------------|-------------|-----------|---------------------|
| | Cooperation | Belief | Cooperation |
| Normalized loss l | -0.004 | -0.003 | -0.002 |
| | (0.024) | (0.010) | (0.023) |
| Normalized gain g | 0.021 | -0.002 | 0.022 |
| | (0.017) | (0.007) | (0.016) |
| Belief | | | 0.512*** (0.077) |
| B_is_Coop | -0.310*** | -0.548*** | -0.029 |
| | (0.034) | (0.015) | (0.053) |
| BeliefThenChoice | -0.035 | 0.005 | -0.037 |
| | (0.034) | (0.014) | (0.033) |
| Age | 0.001 | -0.000 | 0.001 |
| | (0.002) | (0.001) | (0.001) |
| Female | 0.009 | -0.004 | 0.011 |
| | (0.035) | (0.015) | (0.034) |
| Ethnicity: Asian | -0.128 | -0.018 | -0.119 |
| | (0.110) | (0.075) | (0.095) |
| Ethnicity: Black | 0.202*** | 0.058* | 0.172*** |
| | (0.063) | (0.030) | (0.062) |
| Ethnicity: Other | 0.111 | 0.099** | 0.060 |
| | (0.098) | (0.050) | (0.093) |
| Spending/Income: | 0.010 | -0.008 | 0.014 |
| Above median | (0.038) | (0.017) | (0.037) |
| Spending/Income: | 0.112 | -0.008 | 0.116 |
| Prefer not to say | (0.349) | (0.072) | (0.385) |
| Political attitude: | 0.065 | 0.056 | 0.036 |
| Left | (0.074) | (0.044) | (0.066) |
| Political attitude: Right | 0.064 | 0.067 | 0.030 |
| | (0.072) | (0.042) | (0.064) |
| Previous experience in experiments | 0.012 | -0.052*** | 0.038 |
| | (0.035) | (0.015) | (0.034) |
| Constant | 0.536*** | 0.753*** | 0.150 |
| | (0.128) | (0.068) | (0.133) |
| R^2 | 0.12 | 0.64 | 0.16 |
| Obs. | 802 | 802 | 802 |

Notes: Coefficients of a linear probability model (Cols. 1, 3) or linear model (Col. 2) with robust standard errors in parentheses. * p < 0.1; *** p < 0.05; **** p < 0.01.

Table J4. Normalized loss l, normalized gain g, beliefs, and cooperation in the high-EFF games of Study 1.

| | | | Study 1. | | | |
|--|------------------|-------------------|------------------|-------------------|------------------|-------------------|
| Dependent variable: | (1) | (2) | (3) | (4) | (5) | (6) |
| cooperation dummy | Study 1: | Study 1: | Study 1: | Study 1: | Study 1: | Study 1: |
| | UoN | AMT | UoN | AMT | UoN | AMT |
| | Cooperation | Cooperation | Belief | Belief | Cooperation | Cooperation |
| Normalized loss l | -0.106 | -0.163 | -0.046 | 0.052 | -0.075 | -0.184 |
| | (0.134) | (0.140) | (0.083) | (0.089) | (0.123) | (0.145) |
| Normalized gain g | -0.133*** | -0.237*** | -0.029 | -0.078** | -0.114** | -0.204*** |
| 5 · · · · · · · · · · · · · · · · · · · | (0.047) | (0.056) | (0.028) | (0.033) | (0.047) | (0.054) |
| Belief | | | | | 0.668*** | 0.434*** |
| Delici | | | | | (0.071) | (0.092) |
| Round | -0.033*** | -0.025*** | -0.015*** | -0.020*** | -0.022*** | -0.018** |
| Kound | (0.008) | (0.008) | (0.005) | (0.005) | (0.007) | (0.008) |
| D is Coom | 0.035 | -0.054 | -0.045* | -0.174*** | 0.063* | 0.027 |
| B_is_Coop | (0.037) | (0.046) | (0.026) | (0.029) | (0.035) | (0.045) |
| D 1' (TI CI ' | 0.033 | -0.018 | 0.002 | -0.015 | 0.032 | -0.014 |
| BeliefThenChoice | (0.035) | (0.035) | (0.022) | (0.023) | (0.032) | (0.036) |
| | 0.007 | 0.004 | -0.000 | 0.004* | 0.008 | 0.002 |
| Age | (0.007) | (0.003) | (0.003) | (0.002) | (0.006) | (0.003) |
| | -0.007 | 0.084 | 0.002 | 0.059 | -0.008 | 0.059 |
| Female | (0.059) | (0.072) | (0.036) | (0.039) | (0.052) | (0.063) |
| | -0.003 | 0.097 | -0.032 | 0.110 | 0.019 | 0.049 |
| Ethnicity: Asian | (0.070) | (0.107) | (0.042) | (0.074) | (0.059) | (0.097) |
| | -0.009 | 0.103 | 0.033 | 0.034 | -0.030 | 0.088 |
| Ethnicity: Black | (0.092) | (0.121) | (0.041) | (0.064) | (0.082) | (0.102) |
| | 0.011 | 0.043 | 0.045 | -0.015 | -0.019 | 0.049 |
| Ethnicity: Other | (0.085) | (0.149) | (0.046) | (0.081) | (0.075) | (0.136) |
| Business/Eco- | -0.078 | (0.1.5) | -0.056 | (0.001) | -0.041 | (0.120) |
| nomics major | (0.067) | | (0.037) | | (0.059) | |
| , and the second | 0.073 | -0.092 | 0.061 | -0.031 | 0.032 | -0.078 |
| Spending/Income: Above median | (0.062) | (0.076) | (0.041) | (0.041) | (0.052) | (0.068) |
| | | | | ` , | | |
| Spending/Income: Prefer not to say | 0.148 (0.134) | -0.196 (0.225) | 0.011 (0.075) | -0.074 (0.238) | 0.140 (0.102) | -0.166 (0.132) |
| • | | | | | | |
| Political attitude: Left | 0.241*** | -0.115 (0.103) | 0.046 (0.045) | -0.098 (0.064) | 0.210*** | -0.072 (0.088) |
| | (0.075) | (0.103) | | ` / | (0.065) | , |
| Political attitude: | 0.183** | -0.121 | 0.032 | -0.081 | 0.162** | -0.085 |
| Right | (0.087) | (0.102) | (0.050) | (0.059) | (0.077) | (0.087) |
| Previous experience | -0.060 | 0.018 | -0.034 | 0.023 | -0.037 | 0.008 |
| in experiments | (0.056) | (0.076) | (0.032) | (0.043) | (0.048) | (0.067) |
| Constant | 0.354* | 0.845*** | 0.601*** | 0.643*** | -0.049 | 0.567*** |
| | (0.209) | (0.165) | (0.105) | (0.110) | (0.188) | (0.160) |
| Within R^2 | 0.07 | 0.09 | 0.04 | 0.20 | 0.18 | 0.10 |
| Obs. (Clusters) | 616 (154) | 476 (119) | 616 (154) | 476 (119) | 616 (154) | 476 (119) |

Notes: Coefficients of a random effects linear probability model (Cols. 1-2, 5-6) or linear model (Col. 3-4) with robust standard errors clustered on participants in parentheses. * p < 0.1; ** p < 0.05; *** p < 0.01.

Table J5. Normalized loss l, normalized gain g, beliefs, and cooperation in the high-EFF games of Study 2.

| Dependent variable: | (1) | (2) | (3) |
|------------------------------------|-------------|-----------|---------------------|
| | Cooperation | Belief | Cooperation |
| Normalized loss l | -0.010 | -0.104 | 0.052 |
| | (0.142) | (0.064) | (0.138) |
| Normalized gain g | -0.083* | -0.012 | -0.076 |
| | (0.050) | (0.022) | (0.048) |
| Belief | | | 0.604*** (0.075) |
| B_is_Coop | -0.313*** | -0.553*** | 0.020 |
| | (0.034) | (0.015) | (0.054) |
| BeliefThenChoice | -0.034 | 0.025* | -0.049 |
| | (0.033) | (0.015) | (0.032) |
| Age | -0.000 | 0.000 | -0.000 |
| | (0.002) | (0.001) | (0.002) |
| Female | -0.004 | 0.011 | -0.010 |
| | (0.034) | (0.015) | (0.034) |
| Ethnicity: Asian | -0.053 | 0.111** | -0.120 |
| | (0.096) | (0.053) | (0.078) |
| Ethnicity: Black | -0.030 | 0.013 | -0.038 |
| | (0.070) | (0.040) | (0.067) |
| Ethnicity: Other | -0.028 | -0.000 | -0.028 |
| | (0.085) | (0.033) | (0.079) |
| Spending/Income: | -0.049 | -0.022 | -0.035 |
| Above median | (0.037) | (0.016) | (0.036) |
| Spending/Income: | 0.010 | 0.062 | -0.027 |
| Prefer not to say | (0.190) | (0.088) | (0.154) |
| Political attitude: | 0.069 | -0.007 | 0.073 |
| Left | (0.073) | (0.040) | (0.067) |
| Political attitude: Right | 0.003 | -0.001 | 0.004 |
| | (0.072) | (0.038) | (0.066) |
| Previous experience in experiments | 0.062* | 0.016 | 0.052 |
| | (0.035) | (0.015) | (0.034) |
| Constant | 0.791*** | 0.812*** | 0.301** |
| | (0.121) | (0.056) | (0.132) |
| R^2 | 0.12 | 0.65 | 0.18 |
| Obs. | 799 | 799 | 799 |

Notes: Coefficients of a linear probability model (Cols. 1, 3) or linear model (Col. 2) with robust standard errors in parentheses. * p < 0.1; *** p < 0.05; **** p < 0.01.

We define the expected normalized payoff loss from cooperation as $(1 - Belief) \times l + Belief \times g$. This is a composite measure of belief-weighted normalized payoff indices. Its first part captures the normalized loss from cooperating against a defector weighted by the expected likelihood of defection. The second part consists of the foregone normalized gain when cooperating against a cooperator weighted by the expected likelihood of cooperation.

Table J6 reports the results of regression models that include a cooperation dummy as dependent variable and the composite measure as explanatory variable separately for each of the three samples. The control variables are round, order of tasks, order of choices, age, gender, ethnicity, Business/Economics major (UoN only), spending/income, political attitude, and previous experience in experiments. The expected normalized payoff loss from cooperation has a highly significant negative effect on cooperation in each of the three samples. This shows that cooperation behavior is jointly affected by the games' incentives as captured by the normalized indices and expected behavior in others.

Table J6. Expected normalized payoff loss from cooperation.

| Dependent variable: cooperation | (1) Study 1 UoN | (2) Study 1 AMT | (3) Study 2 |
|--|-----------------------|-----------------------|----------------------|
| Expected normalized payoff loss from cooperation | -0.063*** (0.011) | -0.044*** (0.012) | -0.031*** (0.011) |
| Control variables | Yes | Yes | Yes |
| Constant | 0.198 (0.184) | 0.505*** (0.139) | 0.687*** (0.070) |
| (Within) R ² | 0.06 | 0.02 | 0.11 |
| Obs. (Clusters) | 1232 (154) | 952 (119) | 1601 |

Notes: Coefficients of a random effects linear probability model (Col. 1-2) and a linear probability model (Col. 3) with robust standard errors clustered on participants in parentheses. * p < 0.1; ** p < 0.05; *** p < 0.01.

Table J7. K-index, beliefs, and cooperation in Study 1.

| | T tible 97 | · It mach, sen | iers, and ecop | cration in St | ady 1. | |
|----------------------------------|-------------|----------------|-------------------|---------------|--------------|--------------|
| Dependent variable: | (1) | (2) | (3) | (4) | (5) | (6) |
| cooperation dummy | Study 1: | Study 1: | Study 1: | Study 1: | Study 1: | Study 1: |
| | UoN | AMT | UoN | AMT | UoN | AMT |
| | Cooperation | Cooperation | Belief | Belief | Cooperation. | Cooperation. |
| K-index | 0.382*** | 0.297*** | 0.199*** | 0.126*** | 0.265*** | 0.242*** |
| K-macx | (0.073) | (0.084) | (0.041) | (0.049) | (0.072) | (0.083) |
| D-1:-£ | | | | | 0.583*** | 0.440*** |
| Belief | | | | | (0.061) | (0.068) |
| | -0.023*** | -0.012** | -0.015*** | -0.011*** | -0.014*** | -0.007 |
| Round | (0.006) | (0.006) | (0.003) | (0.004) | (0.005) | (0.006) |
| | 0.006 | -0.025 | -0.065*** | -0.158*** | 0.044* | 0.049 |
| B_is_Coop | (0.024) | (0.033) | (0.019) | (0.025) | (0.026) | (0.031) |
| | | | | 0.004 | ` ' | ` ′ |
| BeliefThenChoice | 0.030 | -0.029 | -0.005 (0.013) | | 0.033 | -0.031 |
| | (0.023) | (0.027) | (0.013) | (0.016) | (0.021) | (0.026) |
| Age | 0.008 | 0.006* | 0.001 | 0.005** | 0.008 | 0.004 |
| C | (0.007) | (0.003) | (0.003) | (0.002) | (0.006) | (0.003) |
| Female | 0.014 | 0.094 | 0.006 | 0.029 | 0.011 | 0.081 |
| 1 cmare | (0.057) | (0.065) | (0.034) | (0.033) | (0.049) | (0.056) |
| Ethnicity: Asian | 0.037 | 0.093 | -0.024 | 0.077 | 0.051 | 0.059 |
| Euimenty. Asian | (0.064) | (0.099) | (0.038) | (0.068) | (0.054) | (0.085) |
| Ed. 11. DI. I | 0.072 | 0.041 | 0.081* | 0.056 | 0.024 | 0.017 |
| Ethnicity: Black | (0.086) | (0.101) | (0.042) | (0.052) | (0.082) | (0.090) |
| | 0.002 | 0.014 | 0.021 | 0.003 | -0.010 | 0.013 |
| Ethnicity: Other | (0.083) | (0.131) | (0.045) | (0.077) | (0.067) | (0.111) |
| Business/Eco- | -0.101* | , | -0.082** | , | -0.053 | , |
| nomics major | (0.061) | | (0.033) | | (0.053) | |
| - | 0.072 | -0.079 | 0.080** | -0.031 | 0.025 | -0.065 |
| Spending/Income: Above median | (0.072) | (0.064) | (0.039) | (0.036) | (0.049) | (0.055) |
| | | | | | ` ' | ` ′ |
| Spending/Income: | 0.136 | -0.164 | 0.043 | -0.064 | 0.111 | -0.137 |
| Prefer not to say | (0.122) | (0.213) | (0.073) | (0.155) | (0.096) | (0.150) |
| Political attitude: | 0.234*** | -0.128 | 0.055 | -0.078 | 0.202*** | -0.093 |
| Left | (0.065) | (0.095) | (0.042) | (0.061) | (0.056) | (0.078) |
| Political attitude: | 0.180** | -0.103 | 0.047 | -0.085 | 0.152** | -0.066 |
| Right | (0.077) | (0.091) | (0.046) | (0.059) | (0.067) | (0.075) |
| Previous experience | -0.067 | 0.012 | -0.022 | -0.011 | -0.054 | 0.017 |
| in experiments | (0.051) | (0.068) | (0.030) | (0.037) | (0.042) | (0.058) |
| | -0.025 | 0.342** | 0.435*** | 0.482*** | -0.279* | 0.127 |
| Constant | (0.180) | (0.140) | (0.092) | (0.096) | (0.151) | (0.123) |
| Within R ² | 0.06 | 0.03 | 0.07 | 0.13 | 0.14 | 0.06 |
| Obs. (Clusters) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) | 1232 (154) | 952 (119) |
| V. C. C. | 1 232 (137) | 1.1.11 | 1232 (137) | 11(012) | 1232 (137) | 1 1 1 1 |

Notes: Coefficients of a random effects linear probability model (Cols. 1-2, 5-6) or linear model (Col. 3-4) with robust standard errors clustered on participants in parentheses. * p < 0.1; ** p < 0.05; *** p < 0.01.

Table J8. K-index, beliefs, and cooperation in Study 2.

| Dependent variable: | (1) | (2) | (3) |
|------------------------------------|-------------|-----------|---------------------|
| | Cooperation | Belief | Cooperation |
| K-index | 0.228*** | 0.069** | 0.189*** |
| | (0.071) | (0.032) | (0.070) |
| Belief | | | 0.564*** (0.053) |
| B_is_Coop | -0.309*** | -0.550*** | 0.001 |
| | (0.024) | (0.011) | (0.038) |
| BeliefThenChoice | -0.035 | 0.014 | -0.043* |
| | (0.024) | (0.011) | (0.023) |
| Age | 0.001 | 0.000 | 0.001 |
| | (0.001) | (0.000) | (0.001) |
| Female | 0.001 | 0.004 | -0.001 |
| | (0.024) | (0.011) | (0.024) |
| Ethnicity: Asian | -0.073 | 0.067 | -0.111* |
| | (0.073) | (0.045) | (0.060) |
| Ethnicity: Black | 0.087* | 0.035 | 0.067 |
| | (0.048) | (0.025) | (0.046) |
| Ethnicity: Other | 0.041 | 0.049 | 0.014 |
| | (0.065) | (0.030) | (0.061) |
| Spending/Income: | -0.022 | -0.016 | -0.013 |
| Above median | (0.026) | (0.012) | (0.026) |
| Spending/Income: | 0.045 | 0.074 | 0.004 |
| Prefer not to say | (0.158) | (0.063) | (0.147) |
| Political attitude: | 0.064 | 0.019 | 0.053 |
| Left | (0.052) | (0.030) | (0.047) |
| Political attitude: Right | 0.032 | 0.029 | 0.016 |
| | (0.051) | (0.029) | (0.046) |
| Previous experience in experiments | 0.035 | -0.017 | 0.045* |
| | (0.025) | (0.011) | (0.024) |
| Constant | 0.570*** | 0.730*** | 0.158** |
| | (0.073) | (0.038) | (0.078) |
| $\overline{R^2}$ | 0.11 | 0.64 | 0.17 |
| Obs. | 1601 | 1601 | 1601 |

Notes: Coefficients of a linear probability model (Cols. 1, 3) or linear model (Col. 2) with robust standard errors in parentheses. * p < 0.1; *** p < 0.05; **** p < 0.01.

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