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On the use of the terminal-value approach in risk-value models

Gregor Dorfleitner¹

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Abstract

We extend risk-value models for valuing streams of risky cash flows by establishing the well-known concept of terminal value in this context. For a constant growth assumption we are able to derive upper and lower bounds for the terminal value in the case of a translation-invariant and in the case of a position-invariant risk measure. For both cases we also obtain an exact formula under a special growth assumption for the future cash flows. Furthermore, we provide results on the applicability of the general findings for the case that the log-return of the risky investment follows a Brownian motion.

Keywords New approach · Business valuation · Project valuation · Terminal value · Risk measure

1 Introduction

The terminal value concept has a long-standing tradition in equity valuation (see e.g. Penman 1998; Courteau et al. 2001) and in the DCF methods used for company or project valuation (Massari et al. 2016, ch. 11). The general idea of utilizing a terminal value is that in a setting with an infinite time horizon and periodically accruing cash flows (or dividends or profits), one first assumes constant (or no) growth after a certain period. Second, the perpetuity formula is applied to find a finite present value of all those cash flows (dividends, profits) that lie after the certain period. In practical set-ups, the terminal value is used frequently (Friedl and Schwetzler 2011).

As in the traditional methods the risk of the cash flows is accounted for by the level of capital costs, i.e. as part of the discount factor, it is relatively easy to obtain terminal values of the form

$$\frac{CF}{c - g} = \sum_{t=1}^{\infty} \frac{CF(1 + g)^{t-1}}{(1 + c)^t}, \quad (1)$$

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where CF represents the (annual) expected cash flow, c the cost-of-capital rate and g the (yearly) growth factor of the cash flows.

Recently, Dorfleitner and Gleißner (2018) have introduced a new valuation concept, called the risk-value-model valuation¹. This concept is based on certainty equivalents derived from risk measures and accounts for the risk by subtracting a (time dependent) risk premium from the expected cash flow in the numerator. In a second step, the certainty equivalent is discounted with the riskless interest rate. The concept is an appealing alternative to traditional DCF valuation methods, which are still based on the empirically falsified capital asset pricing model. However, due to the special form of the valuation formula of a single cash flow

$$\frac{CF_t - \text{risk}(CF_t) \cdot \lambda_t}{(1+r)^t},$$

where λ_t represents the risk premium of period t and r the riskless interest rate, the formula (1) obviously cannot be applied straight forward.

This article answers the research question under which assumptions it is possible to make use of the popular terminal value concept in the risk-value-model framework. It needs to be stated very clearly that this article is the answer to the search for suitable assumptions for upper or lower bounds or even a closed-form for the terminal value. This answer is given from a mathematical point of view. Thus, the various (and mutually exclusive) assumptions are not generally economically justified by our research, which is impossible as they are properties of the expected values and risk measures of the cash flows to be valued and their validity therefore depends on the concrete valuation problem. Thus, a potential user needs to make sure whether the required assumption for the result he or she wishes to apply can be regarded sensible for the valuation problem in which he or she wants to use the terminal value. Our research implies a large step forward for valuation practitioners, who prefer to model some periods explicitly and then make some simplifying assumptions for the rest. In this regard the model is very useful as it brings the terminal value concept into risk-value models. The (relatively new) valuation concept of risk-value models as such is useful as it enables valuation resting only upon a risk measure and the expected value, i.e. reduces informational prerequisites.

In this paper, we derive upper and lower bounds for the terminal value under a constant growth assumption. Additionally, we formulate the assumption of a special development of risk over time, which even allows us to derive an exact formula for the terminal value. Dorfleitner and Gleißner (2018) present two ways of accounting for multi-periodicity, namely separate valuation and cumulation. We restrict ourselves to the first variant as the cumulation approach of carrying the risk to a final period appears neither to be feasible nor economically sensible for an infinite number of cash flows. The separate valuation decomposes a stream of cash flows into single cash flows, each of which is valued on a one-period basis (over a varying time interval). Those findings of Dorfleitner and Gleißner (2018) on which our considerations are based are comprehensively displayed in the “Appendix”.

The remainder of the paper is structured as follows: In the next section, we derive general results on upper and lower bounds for the terminal value if the cash flows and their risk are expected to grow constantly. We also give an exact formula for the terminal value if the cash flows follow a special growth assumption. Sect. 3 then presents implications of the general results in the case of a Geometric Brownian motion as a model for the price dynamics of the

¹ While the term goes back to the work of Sarin and Weber (1993), the concept itself is different as no utility functions are needed to establish the risk-value-models valuation. Newer contributions on the topic of project selection such as Dixit and Tiwari (2020) can be seen as evidence that risk measures are indeed more and more used in project decisions.

market portfolio. In Sect. 4, we analyze the applicability of our results for the risk measures value-at-risk and standard deviation. The final section concludes with a general discussion.

2 General solutions to the terminal-value problem

To commence our considerations, let us first formulate the very problem that is to be solved. Consider the infinite stream of cash flows X_1, X_2, \dots . If N is the period in which we wish to place the terminal value TV_N , then the following consideration of equal present values of the original stream of cash flows and of the first N cash flows plus a terminal value implicitly defines TV_N through

$$\sum_{t=1}^{\infty} V_0(X_t) = \sum_{t=1}^N V_0(X_t) + \sum_{t=N+1}^{\infty} V_0(X_t) = \sum_{t=1}^N V_0(X_t) + \frac{TV_N}{(1+r)^N},$$

where $V_0(\cdot)$ denotes the valuation of a cash flow at time $t = 0$. Then we have $TV_N = (1+r)^N \cdot \sum_{t=N+1}^{\infty} V_0(X_t)$. Note that the terminal value is not a present value, but a future value at $t = N$. Accordingly, throughout this paper we call N the *valuation period* of the terminal value and $N + 1$ the *starting period*, as X_{N+1} is the first cash flow that maps into the terminal value.

As pointed out above, we next formulate different (and mutually exclusive) assumptions under which we are able to derive results for upper and lower bounds and even a for closed-form solution for the terminal value. These assumptions on the cash flows to be valued are simply stated and not economically justified here. This is because they cannot generally be justified but only in concrete application contexts, dependent on what is known about the expected value and the risk of the cash flows to be valued.

2.1 Output-oriented view with TI risk measure

Let ρ now be a TI risk measure, such as the VaR or the CVaR. In order to be able to find terminal values or at least bounds for it, it is necessary to make assumptions about the behavior of X_t if t goes to infinity.

Assumption 1 (*Constant growth assumption*) Let $E(X_{t+1}) = (1+g)E(X_t)$ and $\rho(X_{t+1}) = (1+g)\rho(X_t)$ for $t \geq N + 1$ with $-1 < g < r$.

Note that if we define X_{t+1} to be equal in distribution to $(1+g)X_t$, the constant growth assumption naturally emerges.² As all relevant risk measures are law invariant (see ‘‘Appendix’’), the constant growth assumption is uncritical from the viewpoint of the risk measure. However, in every application it needs to be clarified whether it is an acceptable model for the stream of cash flows to be valued. With this assumption a first theorem can be stated.

Theorem 1 *If the risk premia $(\lambda_t)_t$ are an increasing function in t with $\lambda_t \leq 1$ for all $t \geq N + 1$ and the cash flows follow the constant growth assumption from $N + 1$ on with $\rho(X_{N+1} - E(X_{N+1})) > 0$, then the terminal value at time $t = N$ has the upper bound*

² This condition is a generalization of the usual growth condition $CF_{t+1} = (1+g)CF_t$ for riskless cash flows or $E(CF_{t+1}) = (1+g)E(CF_t)$ for risky cash flows whose risk is accounted for in the discount factor.

$$\frac{E(X_{N+1})(1 - \lambda_{N+1}) - \rho(X_{N+1})\lambda_{N+1}}{r - g} \tag{2}$$

and the lower bound

$$\frac{-\rho(X_{N+1})}{r - g}. \tag{3}$$

Proof To commence, we note the following two properties, which we need for our further reasoning.

$$(i) \ E(X_{N+t}) = (1 + g)^{t-1}E(X_{N+1}), \ \rho(X_{N+t}) = (1 + g)^{t-1}\rho(X_{N+1})$$

and

$$(ii) \ E(X_{N+1}) - \rho(X_{N+1} - E(X_{N+1}))\lambda_{N+1} \leq \\ \leq E(X_{N+1}) - \rho(X_{N+1} - E(X_{N+1}))\lambda_{N+1}$$

Formally, property (i) can be proven to be a direct consequence of the constant growth assumption with complete induction, while (ii) follows from the monotonicity of $(\lambda_t)_t$ in t and the positivity of the centered risk measure.

Now, let us consider the upper bound. We have:

$$\begin{aligned} TV_N &= \sum_{t=1}^{\infty} \frac{E(X_{N+t}) - \rho(X_{N+t} - E(X_{N+t}))\lambda_{N+t}}{(1+r)^t} \\ &\stackrel{(i)}{=} \sum_{t=1}^{\infty} \frac{(1+g)^{t-1}(E(X_{N+1}) - \rho(X_{N+1} - E(X_{N+1}))\lambda_{N+1})}{(1+r)^t} \\ &\stackrel{(ii)}{\leq} \sum_{t=1}^{\infty} \frac{(1+g)^{t-1}(E(X_{N+1}) - \rho(X_{N+1} - E(X_{N+1}))\lambda_{N+1})}{(1+r)^t} \\ &= (E(X_{N+1}) - \rho(X_{N+1} - E(X_{N+1}))\lambda_{N+1}) \sum_{t=1}^{\infty} \frac{(1+g)^{t-1}}{(1+r)^t} \\ &= \frac{E(X_{N+1})(1 - \lambda_{N+1}) - \rho(X_{N+1})\lambda_{N+1}}{r - g}, \end{aligned}$$

which proves the upper bound claim. To prove the lower bound, we follow a similar reasoning:

$$\begin{aligned} TV_N &= \sum_{t=1}^{\infty} \frac{E(X_{N+t}) - \rho(X_{N+t} - E(X_{N+t}))\lambda_{N+t}}{(1+r)^t} \\ &\geq \sum_{t=1}^{\infty} \frac{E(X_{N+t}) - \rho(X_{N+t} - E(X_{N+t}))}{(1+r)^t} = \sum_{t=1}^{\infty} \frac{-\rho(X_{N+t})}{(1+r)^t} \\ &\stackrel{(i)}{=} -\rho(X_{N+1}) \sum_{t=1}^{\infty} \frac{(1+g)^{t-1}}{(1+r)^t} = \frac{-\rho(X_{N+1})}{r - g}. \end{aligned}$$

This proves the lower bound claim. □

Next, we formulate an assumption under which the terminal value can be determined exactly.

Assumption 2 (*Special growth assumption I*) Let $E(X_{t+1}) = (1 + g)E(X_t)$ for $t \geq N + 1$ with $-1 < g < r$ and let the risk follow special growth according to:

$$\rho(X_{t+1}) = \rho(X_t)(1 + g) \frac{\lambda_t}{\lambda_{t+1}} - E(X_t)(1 + g) \left(1 - \frac{\lambda_t}{\lambda_{t+1}}\right) \quad \text{for } t \geq N + 1. \quad (4)$$

This assumption, which presumably hardly can be justified literally, is purely technical. Note that the risk now scales differently to the expected value, hence there is no equality in distribution between X_{t-1} and a multiple of X_t anymore. Still, the assumption may serve as an approximation to reality in some cases. In any case, under this assumption a closed-form terminal value can be found.

Theorem 2 *If the cash flows X_{N+1}, X_{N+2}, \dots follow the special growth assumption I then the terminal value at time $t = N$ is*

$$TV_N = \frac{E(X_{N+1})(1 - \lambda_{N+1}) - \rho(X_{N+1})\lambda_{N+1}}{r - g} \quad (5)$$

Proof We start by noting some properties, which follow from the special growth assumption I by complete induction. As in the proof of Theorem 1, we have $E(X_{N+t}) = (1 + g)^{t-1}E(X_{N+1})$. For the risk measure, things are more complicated. We can rearrange Eq. (4) and write for $t \geq N + 1$:

$$\rho(X_{t+1}) + E(X_t)(1 + g) = \rho(X_t)(1 + g) \frac{\lambda_t}{\lambda_{t+1}} + E(X_t)(1 + g) \frac{\lambda_t}{\lambda_{t+1}},$$

which is equivalent to

$$\rho(X_{t+1} - E(X_{t+1})) = (1 + g) \frac{\lambda_t}{\lambda_{t+1}} \rho(X_t - E(X_t)).$$

Therefore for an arbitrary $t \geq 1$ we receive:

$$(*) \quad \rho(X_{N+t} - E(X_{N+t})) = (1 + g)^{t-1} \frac{\lambda_{N+1}}{\lambda_{N+t}} \rho(X_{N+1} - E(X_{N+1})).$$

Now we can calculate the terminal value as:

$$\begin{aligned} TV_N &= \sum_{t=1}^{\infty} \frac{E(X_{N+t}) - \rho(X_{N+t} - E(X_{N+t}))\lambda_{N+t}}{(1 + r)^t} \\ &\stackrel{(*)}{=} \sum_{t=1}^{\infty} \frac{(1 + g)^{t-1}E(X_{N+1}) - (1 + g)^{t-1} \frac{\lambda_{N+1}}{\lambda_{N+t}} \rho(X_{N+1} - E(X_{N+1}))\lambda_{N+t}}{(1 + r)^t} \\ &= (E(X_{N+1}) - \rho(X_{N+1} - E(X_{N+1}))\lambda_{N+1}) \sum_{t=1}^{\infty} \frac{(1 + g)^{t-1}}{(1 + r)^t} \\ &= \frac{E(X_{N+1})(1 - \lambda_{N+1}) - \rho(X_{N+1})\lambda_{N+1}}{r - g}, \end{aligned}$$

which proves the claim. □

Note that the exact value under the special growth assumption equals the upper bound under the constant growth assumption, which can be explained by a less aggressive risk development under this assumption.

2.2 Output-oriented view with PI risk measure

Considering a PI risk measure, we begin our considerations with the constant growth assumption (Assumption 1 from above), in which case we can prove a slightly different theorem when compared with the case of a TI risk measure.

Theorem 3 *If the risk premia $(\lambda_t)_t$ are a decreasing function in t with $\lambda_t \geq 0$ for all $t \geq N + 1$ and the cash flows follow the constant growth assumption from $N + 1$ on with $\rho(X_{N+1}) > 0$, then the terminal value at time $t = N$ has the upper bound*

$$\frac{E(X_{N+1})}{r - g} \tag{6}$$

and the lower bound

$$\frac{E(X_{N+1}) - \rho(X_{N+1})\lambda_{N+1}}{r - g} \tag{7}$$

Proof The terminal value can be written as

$$\begin{aligned} TV_N &= \sum_{t=1}^{\infty} \frac{E(X_{N+t}) - \rho(X_{N+t})\lambda_{N+t}}{(1+r)^t} \\ &= \sum_{t=1}^{\infty} \frac{(1+g)^{t-1}E(X_{N+1}) - (1+g)^{t-1}\rho(X_{N+1})\lambda_{N+1}}{(1+r)^t} \end{aligned}$$

From this the upper bound directly follows because we have $\rho(X_{N+1}) > 0$ and $\lambda_t \geq 0$ for all t .

To derive the lower bound, we use the fact that $(\lambda_t)_t$ is a decreasing function in t and thus $\lambda_{N+1} \geq \lambda_{N+t}$ for all t . From this we obtain:

$$\begin{aligned} TV_N &= \sum_{t=1}^{\infty} \frac{(1+g)^{t-1}E(X_{N+1}) - (1+g)^{t-1}\rho(X_{N+1})\lambda_{N+t}}{(1+r)^t} \\ &\geq \sum_{t=1}^{\infty} \frac{(1+g)^{t-1}E(X_{N+1}) - (1+g)^{t-1}\rho(X_{N+1})\lambda_{N+1}}{(1+r)^t} \\ &= (E(X_{N+1}) - \rho(X_{N+1})\lambda_{N+1}) \sum_{t=1}^{\infty} \frac{(1+g)^{t-1}}{(1+r)^t} \\ &= \frac{E(X_{N+1}) - \rho(X_{N+1})\lambda_{N+1}}{r - g} \end{aligned}$$

This proves the lower bound claim. □

Again, we formulate a special growth assumption in this context.

Assumption 3 (*Special growth assumption II*) Let the cash flows for $t \geq N + 1$ fulfill

$$E(X_{t+1}) = (1+g)E(X_t)$$

but the risk follows $\rho(X_{t+1}) = \rho(X_t)(1+g)^{\frac{\lambda_t}{\lambda_{t+1}}}$.

Under this, again rather technical assumption, for which the above remarks apply analogously, a terminal value can be found.

Theorem 4 *If the cash flows X_{N+1}, X_{N+2}, \dots follow the special growth assumption II, then the terminal value at time $t = N$ is*

$$\frac{E(X_{N+1}) - \rho(X_{N+1})\lambda_{N+1}}{r - g} \tag{8}$$

Proof The proof is completely analogous to that of Theorem 2, but with the utilization of the property $\rho(X_{N+t}) = (1 + g)^{t-1} \frac{\lambda_{N+1}}{\lambda_{N+t}} \rho(X_{N+1})$, which follows directly through complete induction from the special growth assumption II. \square

2.3 Input-oriented view

Practically, in the specification of Dorfleitner and Gleißner (2018), λ_t^I exceeds every bound for $t \rightarrow \infty$. Therefore it is sensible to assume that there is some latest period t^* , after which $x_0^{(t)} = 0$ so that $x_0 = \sum_{t=1}^{t^*} x_0^{(t)}$. Note that the problem does not occur in the work of Dorfleitner and Gleißner (2018) as there the number of cash flows to be valued is finite.

The general valuation formula then is

$$\sum_{t=1}^{\infty} \frac{E(X_t) - x_0^{(t)}\lambda_t^I}{(1+r)^t} = \sum_{t=1}^{t^*} \frac{E(X_t) - x_0^{(t)}\lambda_t^I}{(1+r)^t} + \sum_{t=t^*+1}^{\infty} \frac{E(X_t)}{(1+r)^t} \tag{9}$$

Let now $E(X_{t+1}) = (1 + g) \cdot E(X_t)$ for $t \geq N + 1 \geq t^* + 1$ with $-1 < g < r$ (constant growth assumption). Then a terminal can be determined in a straightforward manner.

As $x_0^{(t)} = 0$ for all $t \geq N + 1$, the term $E(X_t) - x_0^{(t)}\lambda_t^I$ equals $E(X_t)$ for $t \geq N + 1$. Then the terminal value at time $t = N$ can be written as

$$TV_N = \sum_{t=N+1}^{\infty} \frac{E(X_t)}{(1+r)^{t-N}} = \frac{E(X_{N+1})}{r - g} \tag{10}$$

Note that the formula is a direct application of (1) with $c = r$.

3 A concrete specification

In order to specify a concrete valuation mechanism, we need to choose a model for the market portfolio return dynamics. To this end, we adhere to Dorfleitner and Gleißner (2018) and model the log-return of the market portfolio as a Brownian Motion, which implies a simple scaling of the moments along the time axis. Let μ' and σ be the expected value and the standard deviation of the market portfolio’s one-year log-return R'_M and let r' be the continuously compounded interest with $r' = \ln(1 + r)$ and q_p be the p -quantile of the standard normal distribution. Note that we abstain from using the subindex M for the ease of notation. Let us further assume that $\mu' > r'$, which is the common assumption for the return of a risky investment. Then $t\mu', \sqrt{t}\sigma$ and tr' are the corresponding values for a t -year period.

We restrict the following analysis to the value-at-risk as TI risk measure and the standard deviation as the PI risk measure. Although these risk measures have their limitations and shortcomings, they still can be regarded as very popular in practice as well as in theoretical modeling. Furthermore, a similar analysis as carried out below could of course be done for

Table 1 Overview of employed mathematical notation for the concrete specification

continuously compounded interest rate	r'
p -quantile of the standard normal distribution	q_p
one-year log-return of the market portfolio	R'_M
expected value of R'_M	μ'
standard deviation of R'_M	σ
expected excess return of the market (over the riskless interest rate)	π

other risk measures such as the conditional value-at-risk. In this regard, the following is just a first step of what is possible.

Table 1 contains a brief overview of the special notation employed in this subsection.

3.1 Risk premium sequences

The formula for λ_t with ρ being the (TI) value-at-risk at confidence level $1 - p$ (with $p < 0.5$) then is:

$$\lambda_t^{VaR} := \frac{e^{t\mu' + t\frac{\sigma^2}{2}} - e^{tr'}}{e^{t\mu' + t\frac{\sigma^2}{2}} - e^{t\mu' + q_p\sqrt{t}\sigma}}. \tag{11}$$

The proof has already been provided by Dorfleitner and Gleißner (2018).

If we specify the risk measure ρ to be the (PI) standard deviation, then λ_t , according to (24), can be expressed as:

$$\lambda_t^{SD} := \frac{e^{t\mu' + t\frac{\sigma^2}{2}} - e^{tr'}}{e^{t\mu' + t\frac{\sigma^2}{2}} \sqrt{e^{t\sigma^2} - 1}}. \tag{12}$$

The numerator is derived identically to that of (11), while the denominator emerges from the variance formula for log-normal distributions (Johnson et al. 1994, ch. 14).

Example To illustrate the properties of the two risk premium sequences $(\lambda_t^{VaR})_t$ and $(\lambda_t^{SD})_t$, we analyze the behavior over time for a concrete very realistic parameter set, namely $\mu' = 0.06$, $r' = 0.02$, $\sigma = 0.22$ and $p = 0.5\%$, i.e. $q_p = -2.576$. The parameters μ' and σ can be considered typical for a broad stock index (see e.g. Berk and DeMarzo 2016, ch. 10), while $r' = 0.02$ represents a rather low long-run interest rate, which is however realistic for recent years. Figure 1 provides the graph of the continuous functions over time ($t > 0$).

3.2 General results

In order to shorten the notation in the remainder of this section, we introduce the expected excess return (over the riskless interest rate) of the market $\pi := \mu' - r'$, which is positive according to the above assumption. Additionally, we drop the sub-index of q_p and simply write q . Note that $q < 0$ because of $p < 0.5$.

Theorem 5 *The continuous function $t \mapsto \lambda_t^{VaR}$ defined on $(0, \infty)$ starts with a value of zero in $t = 0+$, then rises and has a value of one at $t^* = \frac{q^2\sigma^2}{\pi^2}$, is larger than one beyond this point and finally converges to 1 from above for $t \rightarrow \infty$.*

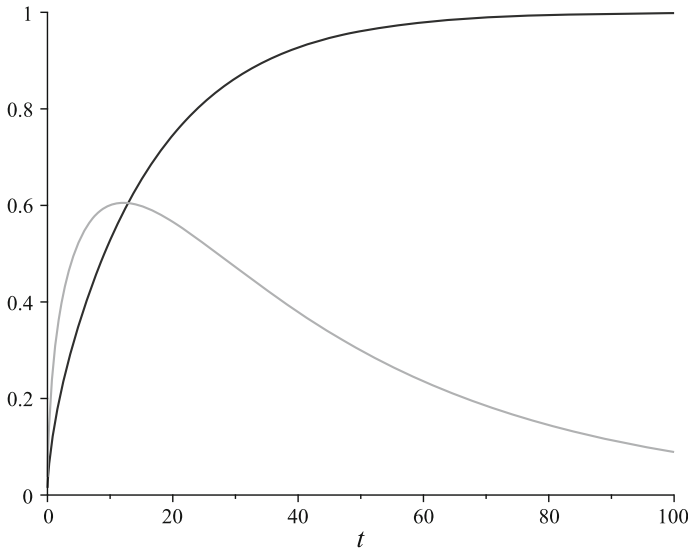


Fig. 1 Graphs of the continuous functions according to λ_t^{VaR} (black line) and λ_t^{SD} (gray line)

Proof We commence with rewriting the risk premium λ_t^{VaR} in the following way:

$$\lambda_t^{VaR} = \frac{e^{t\mu' + t\frac{\sigma^2}{2}} - e^{tr'}}{e^{t\mu' + t\frac{\sigma^2}{2}} - e^{t\mu' + q_p\sqrt{t}\sigma}} = \frac{1 - e^{-t(\pi + \frac{\sigma^2}{2})}}{1 - e^{q\sqrt{t}\sigma - t\frac{\sigma^2}{2}}} \tag{13}$$

As all of the three terms $-(\pi + \frac{\sigma^2}{2})$, q and $-\frac{\sigma^2}{2}$ are negative, both exponential terms in the numerator and the denominator vanish with $t \rightarrow \infty$. Thus $\lambda_t^{VaR} \rightarrow 1$.

We will use the representation (13) for the rest of this proof. From (13) we see that the value of the quotient is equal to one if and only if

$$-t \left(\pi + \frac{\sigma^2}{2} \right) = q\sqrt{t}\sigma - t\frac{\sigma^2}{2} \Leftrightarrow t\pi = -q\sqrt{t}\sigma .$$

As $t > 0$, this is equivalent to $t = \frac{q^2\sigma^2}{\pi^2} =: t^*$. If $t < t^*$, then the numerator of (13) is smaller than the denominator and thus the value of the fraction is less than one (and vice versa for $t > t^*$).

To determine the starting point of the function at $t = 0+$, we need to determine the limit of the fraction according to L'Hospital's rule as both the numerator and the denominator are equal to zero in $t = 0$. Define $u(t) = 1 - e^{-t(\pi + \frac{\sigma^2}{2})}$ and $v(t) = 1 - e^{q\sqrt{t}\sigma - t\frac{\sigma^2}{2}}$. Then (13) equals $u(t)/v(t)$. With

$$u'(t) = \left(\pi + \frac{\sigma^2}{2} \right) e^{-t(\pi + \frac{\sigma^2}{2})}$$

and

$$v'(t) = - \left(\frac{q\sigma}{2\sqrt{t}} - \frac{\sigma^2}{2} \right) e^{q\sqrt{t}\sigma - t\frac{\sigma^2}{2}} ,$$

we have:

$$\begin{aligned} \lim_{t \rightarrow 0+} \frac{u(t)}{v(t)} &= \lim_{t \rightarrow 0+} \frac{u'(t)}{v'(t)} = \lim_{t \rightarrow 0+} \frac{(\pi + \frac{\sigma^2}{2})e^{-t(\pi - \frac{\sigma^2}{2})}}{-\left(\frac{q\sigma}{2\sqrt{t}} - \frac{\sigma^2}{2}\right)e^{q\sqrt{t}\sigma - t\frac{\sigma^2}{2}}} \\ &= \lim_{t \rightarrow 0+} \frac{\sqrt{t}(\pi + \frac{\sigma^2}{2})e^{-t(\pi - \frac{\sigma^2}{2})}}{-\left(\frac{q\sigma}{2} - \sqrt{t}\frac{\sigma^2}{2}\right)e^{q\sqrt{t}\sigma - t\frac{\sigma^2}{2}}} = \frac{0}{-\frac{q\sigma}{2}} = 0. \end{aligned}$$

□

Note that a mathematical treatment of the location and the uniqueness of the maximum and the exact slope behaviour of (13) on $(0, \infty)$ is, on the one hand, not trivial and would require a lot of space. On the other hand, though, it is not really necessary as in the application the function $t \mapsto \lambda_t^{VAR}$ is only evaluated for $t \in \{1, 2, \dots\}$. Therefore, it appears more appropriate to merely prove the rough development over time, as is done in Theorem 5, and leave a more detailed analysis, which will be required as a preparatory step when tackling a real valuation problem, to numerics. We will add more considerations in the next subsection. Note that in any case, Theorem 1 cannot be applied in a straightforward manner because λ_t^{VAR} is not generally smaller than 1 (nor strictly increasing in t) as after t^* it takes values above 1. However, as we will see in the next subsection, Theorem 1 can still be used as a suitable approximation to the problem.

Next, we clarify the rough development over time of λ_t^{SD} .

Theorem 6 *The continuous function according to $t \mapsto \lambda_t^{SD}$ defined on $(0, \infty)$ is strictly positive, it starts at a value of 0 with positively infinite slope in $t = 0+$, has a negative slope for every*

$$t > t^\diamond := \frac{-\ln \frac{\sigma^2}{\sigma^2 + \pi}}{\frac{\sigma^2}{2} + \pi}$$

and converges to 0 for $t \rightarrow \infty$ from above.

Proof We commence with rewriting the risk premium (12) in the following way:

$$\lambda_t^{SD} = \frac{e^{t\mu + t\frac{\sigma^2}{2}} - e^{tr'}}{e^{t\mu + t\frac{\sigma^2}{2}}\sqrt{e^{t\sigma^2} - 1}} = \frac{1 - e^{-t(\pi + \frac{\sigma^2}{2})}}{\sqrt{e^{t\sigma^2} - 1}}. \tag{14}$$

Now, obviously, $e^{-t(\pi + \frac{\sigma^2}{2})}$ is less than one for $t > 0$ and converges to 0 for $t \rightarrow \infty$, while $e^{t\sigma^2} - 1 > 0$. Thus, (14) has a positive value. Concerning the limit behavior, we observe that for $t \rightarrow \infty$ the numerator converges to 1, while the denominator exceeds every bound. Therefore, $\lambda_t \rightarrow 0$.

To determine the starting point of the function at $t = 0+$, we need to determine the limit of the fraction according to L'Hospitale's rule as both the numerator and the denominator are equal to zero in $t = 0$. We define $u(t) = 1 - e^{t(r' - \mu - \frac{\sigma^2}{2})}$ and $v(t) = \sqrt{e^{t\sigma^2} - 1}$. Then (14) equals $u(t)/v(t)$. With $u'(t) = (\pi + \frac{\sigma^2}{2})e^{-t(\pi + \frac{\sigma^2}{2})}$ and $v'(t) = \frac{e^{t\sigma^2} \frac{\sigma^2}{2}}{\sqrt{e^{t\sigma^2} - 1}}$, we have:

$$\lim_{t \rightarrow 0+} \frac{u(t)}{v(t)} = \lim_{t \rightarrow 0+} \frac{u'(t)}{v'(t)} = \lim_{t \rightarrow 0+} \frac{(\pi + \frac{\sigma^2}{2})e^{-t(\pi + \frac{\sigma^2}{2})}}{\frac{e^{t\sigma^2} \frac{\sigma^2}{2}}{\sqrt{e^{t\sigma^2} - 1}}}$$

$$= \lim_{t \rightarrow 0^+} \frac{2\left(\pi + \frac{\sigma^2}{2}\right)e^{-t(\pi+3/2\sigma^2)} \left(\sqrt{e^{t\sigma^2} - 1}\right)}{\sigma^2} = \frac{0}{\sigma^2} = 0$$

To investigate the slope claims, we consider the first derivative according to the quotient rule. First we calculate $u'(t)v(t) - u(t)v'(t)$ as

$$\left(\pi + \frac{\sigma^2}{2}\right)e^{-t(\pi+\frac{\sigma^2}{2})}\sqrt{e^{t\sigma^2} - 1} - \left(1 - e^{-t(\pi+\frac{\sigma^2}{2})}\right)\frac{e^{t\sigma^2}\frac{\sigma^2}{2}}{\sqrt{e^{t\sigma^2} - 1}}. \tag{15}$$

Next, we multiply (15) with $\sqrt{e^{t\sigma^2} - 1}$, which corresponds to expanding the first derivative of $u(t)/v(t)$ with this term. This yields:

$$\begin{aligned} &\left(\pi + \frac{\sigma^2}{2}\right)e^{-t(\pi+\frac{\sigma^2}{2})}\left(e^{t\sigma^2} - 1\right) - \left(1 - e^{-t(\pi+\frac{\sigma^2}{2})}\right)e^{t\sigma^2}\frac{\sigma^2}{2} = \\ &= \left(\pi + \sigma^2\right)e^{-t(\pi-\frac{\sigma^2}{2})} - \left(\pi + \frac{\sigma^2}{2}\right)e^{-t(\pi+\frac{\sigma^2}{2})} - \frac{\sigma^2}{2}e^{t\sigma^2} \end{aligned} \tag{16}$$

To determine the limit of the first derivative at $t = 0+$, we observe that the numerator (16) and the denominator $(e^{t\sigma^2} - 1)^{3/2}$ both vanish at zero and therefore apply L'Hospital's rule. We have

$$\begin{aligned} \lim_{t \rightarrow 0^+} \frac{d}{dt} \frac{u(t)}{v(t)} &= \lim_{t \rightarrow 0^+} \frac{\left(\pi + \sigma^2\right)e^{-t(\pi-\frac{\sigma^2}{2})} - \left(\pi + \frac{\sigma^2}{2}\right)e^{-t(\pi+\frac{\sigma^2}{2})} - \frac{\sigma^2}{2}e^{t\sigma^2}}{\left(e^{t\sigma^2} - 1\right)^{3/2}} \\ &= \lim_{t \rightarrow 0^+} \frac{-\left(\pi + \sigma^2\right)\left(\pi - \frac{\sigma^2}{2}\right)e^{-t(\pi-\frac{\sigma^2}{2})} + \left(\pi + \frac{\sigma^2}{2}\right)\left(\pi + \frac{\sigma^2}{2}\right)e^{-t(\pi+\frac{\sigma^2}{2})} - \frac{\sigma^4}{2}e^{t\sigma^2}}{3\frac{\sigma^2}{2}\sqrt{e^{t\sigma^2} - 1}e^{t\sigma^2}} \\ &= \lim_{t \rightarrow 0^+} \frac{\pi\left(1 + \frac{\sigma^2}{2}\right)}{3\sqrt{e^{t\sigma^2} - 1}} = +\infty. \end{aligned} \tag{17}$$

Now let us assume that $t > t^\diamond$. This is equivalent to

$$-t\left(\pi + \frac{\sigma^2}{2}\right) < \ln \frac{\frac{\sigma^2}{2}}{\sigma^2 + \pi} \Leftrightarrow e^{-t(\pi+\frac{\sigma^2}{2})} < \frac{\frac{\sigma^2}{2}}{\sigma^2 + \pi} \Leftrightarrow \left(\pi + \sigma^2\right)e^{-t(\pi+\frac{\sigma^2}{2})} < \frac{\sigma^2}{2}.$$

With this, (16) can be written as $e^{t\sigma^2} \left(\left(\pi + \sigma^2\right)e^{-t(\pi+\frac{\sigma^2}{2})} - \frac{\sigma^2}{2} \right) - \left(\pi + \frac{\sigma^2}{2}\right)e^{-t(\pi+\frac{\sigma^2}{2})}$, which is obviously less than zero. Therefore, the sign of the derivative of $t \mapsto \lambda_t$ is negative for all t above the positive figure t^\diamond . \square

Note that after starting with a positive slope, the function necessarily possesses a local maximum somewhere between 0 and t^\diamond . Again, we abstain from searching the exact locations of the maxima as this has to be done numerically as a preparatory step of the valuation procedure.

3.3 Application, numerical examples and results

Let us now apply the general results for the specific interpretations of λ_t introduced in this section. We start with the TI risk measure, i.e. the VaR, and a cash flow X_{N+1} with

$E(X_{N+1}) = 2$ and $\rho(X_{N+1}) = -1$. Note that the latter implies $\rho(X_{N+1} - E(X_{N+1})) = 1$ and therefore still guarantees a value reduction due to risk. Let us further assume that the growth factor is $g = 0.01$.

Theorem 1 *Now reverting to Theorem 1, we have to state that the behavior of λ_t^{VaR} over time is such that the assumptions of the theorem are not valid. However, knowing that λ_t^{VaR} converges to 1, the result on the upper bound can be used from a certain N^* onwards. The result on the lower bound is, strictly speaking, not applicable. However, in reality, for many parameter values the maximum value above 1 is so extremely close to 1 that it still makes no difference numerically.*

Example Let us continue the example from above with $\pi = 0.04$, $r = 0.0202$, $\sigma = 0.22$ and $p = 0.5\%$, i.e. $q_p = -2.576$. For these values, the function λ_t^{VaR} strictly increases on $t \in \{1, 2, \dots, 220\}$ and reaches its maximum at 220 with a value of 1.000000356, which is a relative error of less than one millionth and therefore surely negligible.

Thus we have an (approximate) lower bound of

$$\frac{-\rho(X_{N+1})}{0.0202 - 0.01} = \frac{1}{0.0102} = 98.03.$$

The upper bound according to Theorem 1 for the starting period $N + 1 = 11$ is

$$\frac{2(1 - \lambda_{N+1}^{VaR}) - 1 \cdot \lambda_{N+1}^{VaR}}{0.0202 - 0.01} = 139.83,$$

while for $N = 20$ (50) it takes a value of 120.08 (101.25). A rough approximation of the real terminal value could then be to take the arithmetic mean of both bounds yielding the estimates 118.93 ($N = 10$), 109.05 ($N = 20$) and 99.64 ($N = 50$). The real terminal values, according to deeper numerical analysis³, are 103.88 ($N = 10$), 101.11 ($N = 20$) and 98.47 ($N = 50$). While in the latter case the approximate value is already relatively close, the upper bound is generally not as close to the real value as the lower bound.

Theorem 2 *Theorem 2 can of course simply be applied and yields the value of the upper bound of the above example. The question of interest is, which implications follow for the modelling of the cash flows and whether these are realistic. Starting in $t = N + 1$, the model for the risk of the future cash flows is*

$$\begin{aligned} \rho(X_{N+t}) &= 1.01^{t-1} \cdot \frac{\lambda_{N+1}^{VaR}}{\lambda_{N+t}^{VaR}} (\rho(X_{N+1}) + E(X_{N+1})) \\ &\quad - 1.01^{t-1} \cdot E(X_{N+1}) = 1.01^{t-1} \cdot \left(\frac{\lambda_{N+1}^{VaR}}{\lambda_{N+t}^{VaR}} \cdot 1 - 2 \right). \end{aligned}$$

Figure 2 illustrates the development of the risk over time under the special growth assumption for different starting periods. It surely will be hard to find application cases in which the risk of the cash flows exactly evolves like that. However, it is not completely impossible and may be justifiable as an approximation to reality in some settings, especially for late starting periods (cf. the gray solid line).

Next, we consider the standard deviation as risk measure and Theorem 3 and Theorem 4. Here, we again assume $g = 0.01$, $E(X_{N+1}) = 2$, but now $\rho(X_{N+1}) = 1$.

³ This analysis is based on the numerical calculation of the sum $(1 + r)^N \cdot \sum_{t=N+1}^L V_0(X_t)$ for a value of L that is so high that increasing L at most changes the numerical value by less than 10^{-7} . Here, a value of $L = 3000$ is sufficient and, thus, practically the proxy for infinity.

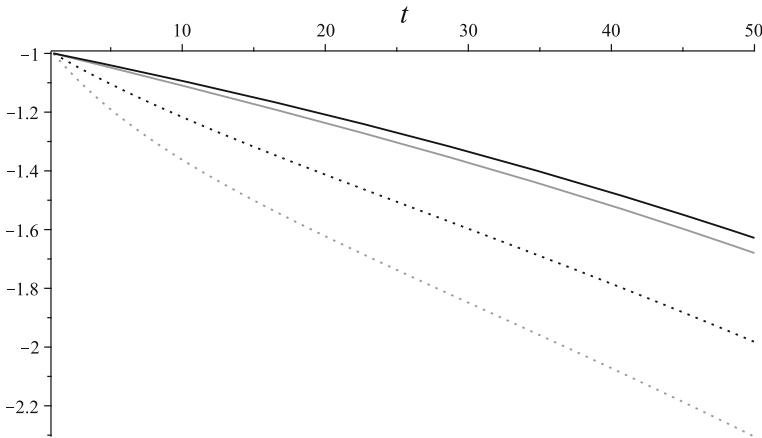


Fig. 2 Graphs of the VaR development according to the special growth assumption dependent on the number of periods after valuation period N for $N = 10$ (gray dotted line), $N = 20$ (black dotted line) and $N = 50$ (gray solid line). For comparison, the black solid line displays the exponential growth development with $g = 0.01$, i.e. the constant growth assumption

Theorem 3 We can state that the behavior of λ_t^{SD} over time is such that the assumptions of Theorem 3 are fulfilled, at least in any case in which the starting period of the terminal value $N + 1$ lies right of t^\diamond .

Example Let us continue the example from above with the same parameter values. For these values the function λ_t^{SD} strictly increases on $t \in \{1, 2, \dots, 12\}$ and reaches its maximum at 12 with a value of 0.6053. After this instant of time, λ_t^{SD} decreases against 0. From the fact that t^\diamond is 22.18, we can derive the information that $t \geq 23$ is a sufficient but not necessary condition for a decreasing λ_t^{SD} function. The upper bound according to Theorem 3 is $\frac{2}{0.0202} = 196.05$.

The lower bound for $N + 1 = 12$ is

$$\frac{E(X_{N+1}) - \rho(X_{N+1}) \lambda_{N+1}^{SD}}{0.0202} = 136.71.$$

while for $N = 20$ and $N = 50$ it takes the values 140.58 and 166.66, respectively. A rough approximation of the real terminal value could then be to take the arithmetic mean of both bounds yielding the estimates 166.38 ($N = 12$), 168.32 ($N = 20$) and 181.36 ($N = 50$). The real terminal values are 175.66 ($N = 12$), 178.78 ($N = 20$) and 187.54 ($N = 50$). In these cases, the rough estimates are relatively close to the real values.

Theorem 4 Again, the Theorem 4 can be applied without any problems and yields the lower bound of the above example. However, the special growth assumption II implies a risk structure over time as follows. Starting at $t = N + 1$, the model for the risk of the futures cash flows must follow

$$\rho(X_{N+t}) = 1.01^{t-1} \cdot \frac{\lambda_{N+1}^{SD}}{\lambda_{N+t}^{SD}} \rho(X_{N+1}) = 1.01^{t-1} \cdot \frac{\lambda_{N+1}^{SD}}{\lambda_{N+t}^{SD}}.$$

Dependent on the starting period of the terminal value $N + 1$, this can have severe consequences. Figure 3 depicts the development of the risk over time implied by the special growth

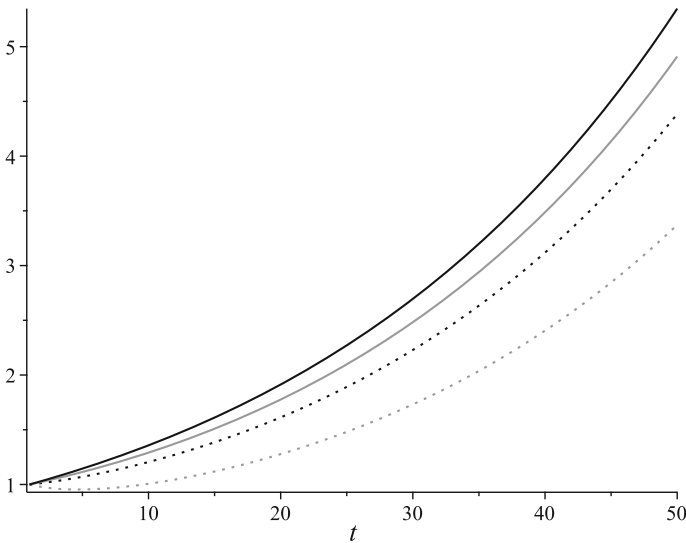


Fig. 3 Graphs of the standard deviation development according to the special growth assumption II dependent on number of periods after valuation period N for $N = 5$ (gray dotted line), $N = 12$ (black dotted line), $N = 20$ (gray solid line) and $N = 50$ (black solid line)

assumption II for several starting periods. Note that risk increases over time much faster than with growth rate $g = 0.01$.⁴ Thus, such a risk development may be regarded realistic only in rare cases. Additionally, a risk evolution as displayed by the gray dotted line (first decreasing, then increasing risk) appears to be extremely specific.

4 Further numerical results and practical considerations

In order to abstract from the concrete example, we conduct a more detailed analysis on the location of the maxima of both risk premium sequences. It is necessary to know about these maxima to be able to apply Theorem 1 and Theorem 3, which build on monotonicity beyond a certain instant of time. We present the results of a systematic numerical investigation in Tables 2 and 3. To this end, we vary the market risk premium π from 0.02 to 0.06 and the volatility of the market from at least 0.14 to at most 0.30 to represent the range of realistic values that could result from statistical estimations or future scenario analysis. In the case of λ_t^{VaR} , we also vary p from 0.1% to 1%. Note that in the case of a real valuation problem, we also need to specify the interest rate r . However, for a calculation of the lambdas, the mentioned parameter settings are sufficient. Building on the general curve sketching results of Theorem 5 and 6, we only need to calculate the values of λ_t for discrete instants of time $t = 1, 2, \dots$ and stop as soon as the slope changes.

The most important findings of this analysis are as follows. In Table 2 one can observe the fact that the deviation from 1 at the maximum of $t \mapsto \lambda_t$ is rather negligible. Only in some rather implausible combinations of high market risk premia and low volatility is the deviation larger than 1% (marked by bold digits). The worst case is the parameter setting of $\pi = 0.06$, $\sigma = 0.14$, $p = 0.01$ with a deviation of approximately 3.26%. Altogether, the maximum

⁴ Indeed, $1.01^{50} = 1.64$, while even for $N = 5$ we have a factor of more than three after 50 periods.

Table 2 Discrete instants of time where λ_t^{VaR} is maximal dependent on a broad variety of realistic values of σ and π .

σ	π								
	0.02	0.025	0.03	0.035	0.04	0.045	0.05	0.055	0.06
Panel A: $p = 0.001$									
0.14	510	337	241	183	145	118	99	84	73
0.16	648	424	301	227	178	144	120	101	87
0.18	805	525	371	277	216	174	144	121	104
0.2	909	637	449	334	260	208	171	144	123
0.22	847	761	535	398	308	246	202	169	144
0.24	767	696	618	467	361	288	236	197	167
0.26	696	637	587	545	418	334	273	227	192
0.28	633	584	541	505	473	381	313	260	220
0.3	576	535	500	468	441	404	354	296	250
Panel B: $p = 0.005$									
0.14	368	246	178	137	110	90	76	66	57
0.16	462	305	219	166	132	108	90	77	67
0.18	570	374	266	201	158	128	106	90	78
0.2	693	451	319	239	187	151	125	105	91
0.22	811	538	379	283	220	177	146	122	105
0.24	767	625	445	331	256	205	169	141	121
0.26	696	637	511	383	296	237	194	162	138
0.28	633	584	526	438	340	271	222	185	157
0.3	576	535	500	455	383	308	251	209	177
Panel C: $p = 0.01$									
0.14	308	208	152	118	95	79	67	58	51
0.16	384	255	184	141	112	93	78	67	58
0.18	471	310	222	168	133	109	91	77	67
0.2	570	373	265	200	157	127	105	89	77
0.22	681	443	313	234	183	148	122	103	88
0.24	735	520	366	273	213	171	141	118	101
0.26	696	589	424	316	245	196	161	135	115
0.28	633	584	479	362	280	224	183	153	130
0.3	576	535	485	410	318	254	207	173	147

The figures in bold digits represent those cases where the maximum value deviates more than 1% from the value of 1. In all other cases the deviation is less than 1%

lambda value is reached relatively late, so that the assumption of strictly increasing λ_t can be maintained as an approximation to reality for every starting period. Note that the above example can be found in Panel B in the middle of the table.

The case of λ_t^{SD} is presented in Table 3. Here we can observe relatively low values of maximizing t . Thus Theorem 3 can be applied for starting periods in the range of around 10 to 20. Also, Theorem 4 can be applied with relatively early starting periods without exhibiting the—possibly implausible—behavior of first decreasing and then increasing risk (cf. red line

Table 3 Discrete instants of time where λ_t^{SD} is maximal dependent on a broad variety of realistic values of σ and π

σ	π								
	0.02	0.025	0.03	0.035	0.04	0.045	0.05	0.055	0.06
0.14	27	25	22	21	19	18	17	16	15
0.16	23	21	20	18	17	16	15	14	13
0.18	20	19	17	16	15	14	14	13	12
0.2	17	16	15	14	14	13	12	12	11
0.22	15	14	13	13	12	12	11	11	10
0.24	13	13	12	11	11	10	10	10	9
0.26	12	11	11	10	10	9	9	9	8
0.28	10	10	9	9	9	9	8	8	8
0.3	9	9	9	8	8	8	7	7	7

in Fig. 3). Such behavior only occurs if the starting period lies before the maximizing t of Table 3.

5 Discussion and conclusion

In this paper, we extend risk-value models for valuing streams of risky cash flows by introducing the concept of terminal value to this framework. As the case of the input-oriented perspective is treatable in a simple manner, the main focus is on the output-oriented view.

For a constant growth assumption, upper and lower bounds for the terminal value in the case of a translation-invariant and in the case of a position-invariant risk measure can be derived. For both cases, we also derive an exact formula, however only under a special growth assumption for the future cash flows, which can be characterized as unrealistic for many application cases, especially in case of the standard deviation being the risk measure. However, the aim of this article is not to provide an economic justification for the assumptions but rather to clarify which assumptions are required to be able to use terminal values or at least upper and lower bounds for these. Furthermore, we demonstrate how the general findings can be applied under the assumption that the price of the market portfolio follows a Geometric Brownian motion. It becomes apparent that in the value-at-risk case under this assumption the prerequisites of the lower bound result are not fulfilled literally. However, for realistic parameter values the corresponding theorem can still serve as an approximation.

As a concrete consequence for applying Theorems 1 to 4, it appears to be advisable to apply these for rather late starting periods $N + 1$, in which case the arithmetic mean of the upper and the lower bound appears to be a relatively good approximation of the correct value. If the cash flow modeling requires relatively early terminal values (say: at $N = 5$), then one can simply value the first periods after the starting period separately (say: from 6 to 30) and only use the terminal value formulae afterwards (say: for $N = 30$). Generally, the upper and lower bound results may be more beneficial for realistic cash flow modeling situations than the closed-form exact terminal value due to the corresponding required assumptions.

Summarizing, the concept of terminal value can in principal be used in risk value models, but only with certain caveats and limitations. This paper therefore marks a significant progress compared with the original framework of Dorfleitner and Gleißner (2018), in which a final

cash flow at some period is required. Generally, it should be noted that as in usual DCF models the value of the stream of cash flows is very sensitive to the specification of the growth rate (Friedl and Schwetzler 2011).

Several generalizations of the concept could be investigated in the future. One idea is to overcome the Brownian motion hypothesis to calculate explicit risk premium sequences. Other dynamics such as dynamics with jumps or long-time mean-reverting processes for the stock price process are also conceivable. Moreover, as with this research risk-value models are also established over an infinite time horizon, it may be an interesting task to investigate whether dynamic (i.e. multi-period) risk measures as introduced by Riedel (2004) can be utilized for a generalisation of the imperfect replication approach. However, it is unclear at the moment how the initial approach of Dorfleitner and Gleißner (2018) is to be generalized that this becomes possible.

In any case, the results found in this paper can be very useful for practitioners, who prefer to model some periods explicitly and then make some simplifying assumptions for the rest. In this regard our contribution extends this concept to risk-value models or at least shows under which assumptions it can be used.

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Appendix A: A short recapitulation of risk-value-model valuation

A.1 Basics

Risk-value model valuation according to Dorfleitner and Gleißner (2018) essentially builds upon a risk measure to determine and quantify the notion of risk and scales the reference investment in a way that it has the same risk as the cash flow to be valued. Thus, we shortly state some necessary essentials about risk measures, which can generally have several properties. For risk-value models three properties are of importance. Let \mathcal{X} be the set of all real-valued random variables on the probability space $(\Omega, \mathfrak{E}, \mathcal{P})$ with existing finite values of $\rho(X)$ for $X \in \mathcal{X}$. Then we define:

- *positive homogeneity* (PH) through $\rho(cX) = c\rho(X)$ for all $c \geq 0$, $X \in \mathcal{X}$,
- *translation invariance* (TI) through $\rho(X + c) = \rho(X) - c$ for all $c \in \mathbb{R}$, $X \in \mathcal{X}$ and
- *position invariance* (PI) through $\rho(X + c) = \rho(X)$ for all $c \in \mathbb{R}$, $X \in \mathcal{X}$.

The prerequisite for risk-value model valuation is the use of a PH risk measure that is simultaneously either TI or PI. Therefore, the expected shortfall, also called conditional value-at-risk, (see e.g. Acerbi and Tasche 2002; Inui and Kijima 2005; Bamberg and Neuhierl 2010) and the value-at-risk are suited as they both are PH and TI. Moreover, PH risk measures fulfilling the PI property such as the standard deviation or TI-variants of PI risk measures, i.e. risk measures applied to $X - E(X)$, which are also deviation measures (Righi 2019), can serve as basis of risk-value-model valuation.

Table 4 Overview of employed general mathematical notation

risk measure	ρ
riskless interest rate	r
risk premium for period t	λ_t
cashflow in period t	X_t
equity capital need for period t	$x_0^{(t)}$
market return over $[0, t]$	$R_{t,M}$
expected market return over $[0, t]$	$\mu_{t,M}$

Table 4 overviews the most important symbols used in this paper.

The general valuation problem can be stated as finding a joint present value of risky future cash flows X_1, X_2, \dots, X_N . Dorfleitner and Gleißner (2018) suggest two generally interesting valuation approaches, the output and the input-oriented view. For the remainder of this “Appendix” we simply state some main content from Dorfleitner and Gleißner (2018) in a nutshell.

A.2 Output-oriented view

For a TI or PI risk measure, the output-oriented method valuation according to Dorfleitner and Gleißner (2018) in a one-period setting (risky cashflow X being realized in one period $t = 1$) starts with the two replication equations

$$E(X) = E(y(1+r) + z(1+R_M)) = y(1+r) + z(1+\mu_M) \quad \text{and} \quad (18)$$

$$\rho(X) = \rho(y(1+r) + z(1+R_M)) \quad , \quad (19)$$

where y is the amount invested risklessly and z the amount invested into the market portfolio. This approach comes to the valuation

$$V(X) = \frac{E(X) - \rho(X - E(X)) \frac{\mu_M - r}{\rho(R_M - \mu_M)}}{1+r} \quad . \quad (20)$$

The valuation of N subsequent joint cashflows X_1, \dots, X_N then is carried out through applying value additivity on the N cashflows by

$$\sum_{t=1}^N \frac{E(X_t) - \rho(X_t - E(X_t)) \lambda_t}{(1+r)^t} \quad , \quad (21)$$

where

$$\lambda_t = \frac{\mu_{t,M} - ((1+r)^t - 1)}{\rho(R_{t,M} - \mu_{t,M})} \quad . \quad (22)$$

Note that in the original framework the cash flow X_N is the last cash flow. However, theoretically also an infinite number of cash flows can be considered with the same model and, if the limit exists, also valued through:

$$\sum_{t=1}^{\infty} \frac{E(X_t) - \rho(X_t - E(X_t)) \lambda_t}{(1+r)^t} \quad .$$

Note that for a PI risk measure, the valuation procedure simplifies to

$$\sum_{t=1}^{\infty} \frac{E(X_t) - \rho(X_t) \lambda_t}{(1+r)^t} \tag{23}$$

with

$$\lambda_t = \frac{\mu_{t,M} - ((1+r)^t - 1)}{\rho(R_{t,M})}. \tag{24}$$

A.3 Input-oriented view

The one-period setup of the input-oriented view with financing restrictions rests on the three equations:

$$x_0 = y + z \tag{25}$$

$$0 = \rho(X) = \rho(y(1+R_D) + z(1+R_M)) \tag{26}$$

$$E(X) = E(v(1+r) + y(1+R_D) + z(1+R_M)), \tag{27}$$

where X is now the cash flow to equity and x_0 represents the initial equity capital need, while the risk measure $\rho(\cdot)$ now is the VaR at the $1 - p$ confidence level and R_D represents the (risky) interest rate for a loan with default probability p . The valuation formula derived is

$$V(X) = \frac{E(X) - x_0 \frac{(1+r_p)(\mu_M - \mu_D) + \rho(R_M - r_p)(\mu_D - r)}{\rho(R_M - r_p)}}{1+r}, \tag{28}$$

where $\mu_D = E(R_D)$ and r_p is the demanded interest rate at default probability p (i.e. the maximum value of R_D). Typically x_0 is larger than or equal to zero. However, if $x_0 < 0$, then the value of X is given by $V(X) = \frac{E(X)}{1+r}$. Whatever the sign of x_0 is, the net present value is obtained as $V(X) - x_0$.

The valuation of N subsequent joint cashflows X_1, \dots, X_N then is carried out through applying value additivity on the N cashflows by

$$\sum_{t=1}^N \frac{E(X_t) - x_0^{(t)} \lambda_t^I}{(1+r)^t} \tag{29}$$

where λ_t^I is a risk premium term, essentially corresponding to the factor multiplied with x_0 in (28) in an annualized version. We do not display λ_t^I in any further detail because it is not of relevance for the considerations of this paper. The quantity in the formula that effectively measures the risk is the equity capital need x_0 , which is allocated to the periods subject to the condition

$$x_0 = \sum_{t=1}^N x_0^{(t)}. \tag{30}$$

As for the input-oriented view the allocation problem is to be solved in any case, (30) also true for an infinite number of cash flows. Moreover, λ_t^I can exceed any boundary for increasing t .⁵ However, if we additionally assume that the $x_0^{(t)}$ finally become 0 after a certain

⁵ At least, this is the case in the concrete specification presented in Dorfleitner and Gleißner (2018).

instant of time, one can also generalize this method to an infinite number of cash flows. Note that then practically from a certain period on the cash flows are valued only through their expected value (as $x_0^{(t)} = 0$).

A.4 Risk-measure related considerations

Last, we complement this recapitulation of Dorfleitner and Gleißner (2018) with some own remarks on risk measures.

The answer to the question which (classes of) risk measures are suited for risk-model valuation is relatively simple. First, coherent risk measures (see Artzner et al. 1999) or spectral risk measures (see e.g. Acerbi 2002; Brandtner and Kürsten 2015) fulfill PH and TI. Furthermore, convex risk measures (Föllmer and Schied 2002) are employable as long as they are PH. The property of law invariance (Kusuoka 2001; Frittelli and Rosazza 2005) is formally not necessary, nevertheless it is fulfilled by all known real-world relevant risk measures.

In the mathematical risk measure literature, we also find dynamic risk measures such as coherent ones (see Riedel 2004 or Artzner et al. 2007) or convex ones (Frittelli and Rosazza 2004). The interesting property of such approaches is the decision consistency over time. However, this research is based on an already established valuation methodology by Dorfleitner and Gleißner (2018), which decomposes a stream of cash flows into single cash flows, each of which is valued on one-period basis (over a varying time interval). In this context, the only instant of time in which a decision is made is $t = 0$. Hence, this framework is also the basis of our contribution. This implies that for constructing a terminal value the one-period valuation (over a varying time interval) is a necessary step. Therefore, dynamic risk measures are in our context no applicable tool.

Furthermore, it is noteworthy that the standard deviation and the VaR together with a Brownian motion assumption of the market log-returns lead to two explicit time-dependent formulae for the risk premia (Dorfleitner and Gleißner 2018). However, many different explicit or implicit risk premia over time can be thought of.

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