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Debunking "fake news" on social media: Short-term and longer-term effects of fact checking and media literacy interventions

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**Debunking "Fake News" on Social Media:
Short-Term and Longer-Term Effects of Fact
Checking and Media Literacy Interventions**

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Debunking “fake news” on social media: short-term and longer-term effects of fact checking and media literacy interventions*

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Abstract

We conduct a randomized survey experiment to compare the short- and longer-term effects of fact checking to a brief media literacy intervention. We show that the impact of fact checking is limited to the corrected fake news, whereas media literacy helps to distinguish between false and correct information more generally, both immediately and two weeks after the intervention. A plausible mechanism is that media literacy enables participants to critically evaluate social media postings, while fact checking fails to enhance their skills. Our results promote media literacy as an effective tool to fight fake news, that is cheap, scalable, and easy-to-implement.

JEL Codes: L51, L82, Z18

Keywords: Covid, Facebook, fact checking, fake news, media literacy, misinformation, nutrition, social media, supplements, survey experiment, vaccine

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1. Introduction

The emergence and spread of “fake news” – i.e., false or misleading information presented as news – has led to widespread concerns (e.g., Lazer et al., 2018). Social media like Facebook and Twitter are especially prone catalysts for the evolution of fake news and have consequently come to the fore of public and academic debates. Indeed, recent evidence suggests that 50% of users who see fake news on social media say that they believe them (Allcott and Gentzkow, 2017).

What helps users to distinguish between false and correct information on social media? Policymakers support fact checkers on the one hand, and media literacy initiatives on the other.¹ Independent fact checking organizations complement such campaigns, and large social media have started to implement AI-based procedures that flag suspicious content, too (Allen et al., 2021). Yet, it is unclear whether these remedies function as desired: empirical evidence on the effectiveness of fact checking is mixed (Vraga and Bode, 2017; Jerit and Zhao, 2020), knowledge on the impact of media literacy is scarce (Guess et al., 2020), and a direct comparison of these interventions does not exist at all.

We address this gap with a large-scale randomized survey experiment on the short- and longer-term effects of fact checking and media literacy interventions. In the experiment, we expose participants to false and correct statements on health-related topics – Corona vaccines and nutrition – that we retrieve from Facebook (“fakes” and “facts”). One group of participants receives additional fact checks that debunk some of the fakes explicitly. Another group gets ten “Tips to spot fake news” before exposure to the fakes and facts as a brief media literacy training. Then, we compare the two treatment groups to participants who do not receive an intervention. To study longer-term effects, we survey the same participants around two weeks later in a second, analogous wave of the experiment.

Our results demonstrate that the effectiveness of fact checking tends to be limited to the fakes that are being corrected, whereas media literacy helps to distinguish between fakes and facts more generally, both in the short- and in the longer-run. A plausible explanation is that the media literacy intervention raises participants’ attention and enables them to critically evaluate the postings’ accuracy. Fact checking, in contrast, turns participants into passive recipients of the specific corrections and thus fails to enhance their skills.

Specifically, we consider three main outcomes: the perceived credibility of fakes and facts, factual knowledge on the topics discussed therein, and attitudes towards Corona vaccination and dietary supplements (the fakes on nutrition promote the consumption of needless protein and vitamin preparations). The idea is to study a coherent cognitive chain: Do the interventions reduce the perceived credibility of fakes (but not of facts)? If yes, does that translate into better factual knowledge? If yes, does this entail a change in attitudes?

We find that both interventions reduce the credibility of fakes on Corona vaccines (which are corrected by fact checks) in the short-run, but only the media literacy intervention reduces the credibility of fakes on nutrition (which are not corrected by fact checks), both in the short- and in the longer-run. Moreover, both interventions improve participants’ factual knowledge in the short-, but only the media literacy intervention in the longer-run. Finally, while the

¹See, e.g., <https://digital-strategy.ec.europa.eu/en/library/disinformation-threat-democracy-brochure> and <https://digital-strategy.ec.europa.eu/en/policies/online-disinformation> (Aug 2022) for further information on efforts by the European Union.

media literacy intervention raises participants’ willingness to get vaccinated (or boosted) against Covid-19 in the short- and in the longer-run, fact checking has no such effect. Crucially, neither intervention reduces the credibility of facts or factual knowledge on the topics discussed therein, i.e., participants do not become more skeptical towards social media postings per se. Hence, in an environment where not every posting can be fact checked, media literacy interventions are likely to be more effective than fact checking on average.

Our subgroup analyses reveal that participants who are well informed from the beginning are less likely to benefit from the interventions than participants whose prior beliefs are further away from the truth. In particular, both the fact checking and the media literacy intervention are *more* effective for fakes on Corona vaccines for supporters of the AfD (“Alternative for Germany”, a far-right populist party known for spreading misinformation on Covid-19). For fakes on nutrition, where participants’ beliefs are much more alike, we observe no such effect. Computing persuasion rates à la DellaVigna and Kaplan (2007) shows that this result can just partly be explained by differences in the proportion of participants who are left to be convinced. However, we also provide evidence that AfD supporters are less certain about their prior knowledge than non-AfD supporters, so the former group may also be easier to convince. In contrast to that, we do not find any systematic effect heterogeneity in terms of education, age, social media usage, support of Corona policy measures, or prior knowledge on current events, health, and nutrition.

A plausible mechanism for our main results is that the media literacy intervention raises participants’ attention and enables them to critically evaluate the Facebook postings, while fact checking fails to enhance their skills. To support the plausibility of this explanation, we show that participants who receive the media literacy intervention are more likely to actively search for further information when they respond to our questions than participants who receive the fact checking or no intervention at all. Moreover, media literacy helps participants to better identify untrustworthy elements in fakes and trustworthy elements in facts. Fact checking, in contrast, has no such effect.

While our main analysis illustrates the effectiveness of fact checking and media literacy interventions in an environment where all participants see fakes and facts, it is agnostic about the extent to which the interventions are able to reverse the harm the fakes are causing. To better interpret the magnitude of our coefficients in that regard, we also compare the three main treatment groups to participants who do not see any Facebook postings at all. We find that exposure to fake news substantially impairs participants’ factual knowledge, and that neither the fact checking nor the media literacy intervention can fully offset the effect. Participants’ attitudes on Corona vaccination and dietary supplements, in contrast, are hardly affected by fakes and especially the media literacy intervention can effectively repeal that impact.

As far as we know, we are the first who pursue a clean comparison of fact checking and media literacy interventions as a means to debunk fake news, whereby we provide a valuable contribution to public and academic debates. Since public resources to combat fake news are limited, it is of utmost importance to understand when and why which remedies are most effective, so that time, money, and effort can be efficiently allocated. Pennycook and Rand (2021), for instance, stress that professional fact-checking is “simply not scalable” (p.396), as it requires substantial time and effort to examine a particular claim, and even if the claim is eventually tagged as false, the warning is likely to be missing during the peak of its spread. Similar caveats apply

to the recent approach of (human) crowd sourced fact-checking (Allen et al., 2021), which is furthermore limited to the linguistic proficiency of crowd workers. AI-based procedures are more efficient than human search, but still in their infancy. Specifically, while they are able to detect potential misinformation, the fact-checking as such still requires human assessment.² We show that in an environment where only a small proportion of fake news can ever be fact checked, media literacy is likely to be more effective than fact checking on average. Moreover, given that displaying a small number of tips and heuristics to users of social media is cheap, scalable, and easy-to-implement, our results promote media literacy interventions as a (potentially more) powerful tool to combat fake news.

Our paper advances the surprisingly small body of research on (digital) media literacy as a means to fight fake news (Guess et al., 2020; Roozenbeek et al., 2022) and adds to a recent literature that acknowledges the limits of fact checking (see Jerit and Zhao, 2020, for a review). E.g., Pennycook and Rand (2019) argue that many users fall for fake news because they fail to reflect; similarly, Pennycook et al. (2020, 2021) show that users share false claims partly because they do not think sufficiently about whether or not the content is accurate. Consistent with what we find, such results advocate media literacy interventions that help users to critically evaluate social media postings as a promising avenue, while assorting fact checking – which fails to enhance users’ skills – as less effective.

The remainder of the paper is organized as follows. Section 2 reviews the related literature on social media, misinformation, and education interventions. Section 3 illustrates the experimental setup and implementation; moreover, we discuss our empirical analysis. Section 4 presents our main results, where we compare the effectiveness of fact checking and media literacy on the credibility of and factual knowledge on fakes, as well as on participants’ attitudes. In Section 5, we show that an increase in attention and the ability to critically evaluate social media postings on behalf of the media literacy intervention is a plausible mechanism for our results. Section 6 presents further results and robustness checks, Section 7 concludes.

2. Related literature

Social media and UGC Our paper is related to two strands of literature. First, it adds to the vibrant and interdisciplinary research on social media and user-generated content (reviewed by Luca, 2015; Zhuravskaya et al., 2020), where it is particularly close to analyses of fake news. This subfield can be further divided into studies on the emergence and spread of fake news (e.g., Allcott and Gentzkow, 2017; Lazer et al., 2018; Guess et al., 2018, 2019; Grinberg et al., 2019; Vosoughi et al., 2018; Bursztyn et al., 2023), and inquiries of potential remedies (reviewed by Lewandowsky et al., 2012; Jerit and Zhao, 2020; Allen et al., 2021). The latter literature focuses on corrective interventions like fact checking: While Bode and Vraga (2015), Vraga and Bode (2017), and Henry et al. (2020), among others, support its effectiveness, other papers find no or even “backfire” effects (e.g., Nyhan and Reifler, 2010, 2015), or they document mixed results, whereby fact checking improves users’ factual knowledge, but struggles to change more deep-rooted perceptions and attitudes (Barrera et al., 2020; Nyhan et al., 2020). Studies on alternative

²See EU Horizon, URL: <https://ec.europa.eu/research-and-innovation/en/horizon-magazine/can-artificial-intelligence-help-end-fake-news> (Aug 2023).

ways to combat fake news are rare. One notable exception are Guess et al. (2020), who assess the effectiveness of Facebook’s “Tips to Spot False News” on discernment between mainstream and false news headlines both among a nationally representative sample in the US and a highly educated online sample in India. Relatedly, Roozenbeek et al. (2022) use five short videos that inoculate people against manipulation techniques commonly used in misinformation and find that they improve manipulation technique recognition, boost confidence in spotting these techniques, increase users’ ability for truth discernment as well as the quality of their sharing decisions.

We contribute to this literature in several ways. First, we pursue a clean comparison of fact checking and media literacy interventions, which has not been done so far. In particular, our experimental setup allows us to study the short- and longer-term effects of fact checking and media literacy interventions in one and the same environment, whereby we can observe when and why which remedy is most effective. In addition, we provide evidence for potential mechanisms behind our results, shifting the research focus from asking *whether* fact checking and media literacy interventions are effective tools to fight fake news to studying *how* they work and *in which case* they fail or succeed.

Most closely related to our study are Barrera et al. (2020) and Guess et al. (2020). Barrera et al. (2020) use a randomized online experiment to expose voters to fakes, facts, and fact checks on immigration in France. Participants are then asked about their posterior beliefs on topics related to immigration, their opinions on immigration policy, as well as their voting intentions. Similar to what we find, Barrera et al. (2020) demonstrate that fake news are highly persuasive, and while fact checking enhances factual knowledge, it fails to offset the fakes’ effect on voting intentions.³ Guess et al. (2020) examine the impact of a digital media literacy intervention on the perceived accuracy of false and correct news headlines and show that participants’ ability for truth discernment increases.

Our results largely confirm these findings, but we extend the preceding analyses in several ways. First, we explore the short- *and* longer-term effects of fact checking *and* media literacy interventions on fakes *and* facts within one experiment, which allows us to directly compare these remedies and draw a sophisticated picture of how and when which type of intervention works. Likewise, we consider a broad range of coherent outcomes – credibility, factual knowledge, and attitudes – and complement our analysis with a thorough examination of potential mechanisms. Finally, we use postings from social media that actually exist and whose content is not necessarily politically loaded, demonstrating that our results hold beyond the partisan context.

We also contribute to a growing body of research arguing that users fall for fake news because they fail to pay sufficient attention (e.g., Pennycook and Rand, 2019; Loewenstein and Wojtowicz, 2023). Pennycook et al. (2020, 2021), for instance, show that users frequently share misinformation because they do not focus on accuracy; politically motivated reasoning, in contrast, seems to play a minor role. Our results support such findings, because we demonstrate that media literacy interventions – which raise users’ attention and help them to actively distinguish between fakes and facts – are on average more effective than just passively receiving fact checks. Moreover, in contrast to previous findings on motivated reasoning (e.g., Lewandowsky et al., 2012; Jerit and Zhao, 2020), we find that our interventions are more effective for supporters of a far-right political party, who are initially much more likely to oppose Corona vaccination. This, too, is consistent

³Similar results are presented by Nyhan et al. (2020).

with the above line of thought, whereby it is often a lack of attention rather than partisanship that causes the ineffectiveness of fact checking.

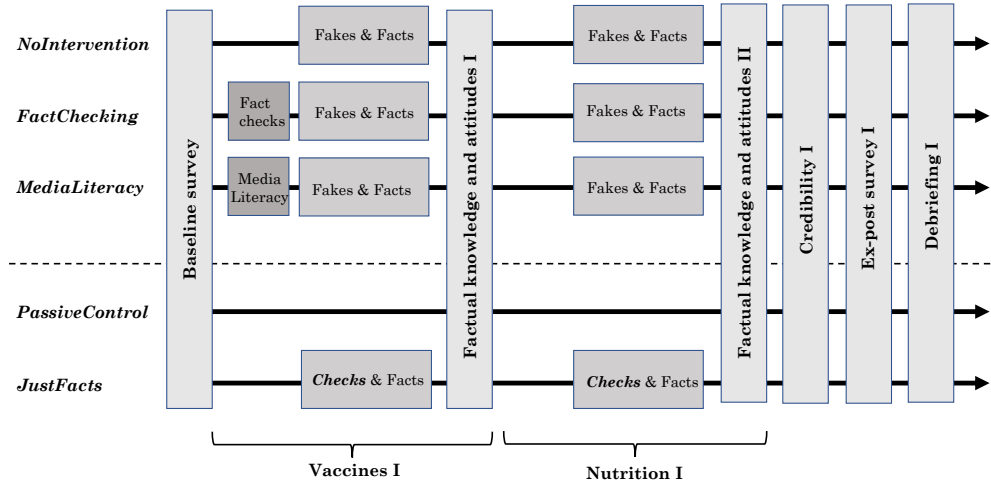
Education interventions Second, our paper is related to the literature on education interventions, most of which focuses on financial literacy. In accordance with our results on the media literacy intervention, a meta-analysis of 76 randomized experiments by Kaiser et al. (2022) reveals that financial education interventions have, on average, positive causal treatment effects on financial knowledge and behavior. The treatment effects are economically meaningful in size and comparable to those realized by education interventions in math and reading (e.g., Hill et al., 2008; Cheung and Slavin, 2016; Fryer Jr, 2017), health (e.g., Rooney and Murray, 1996; Noar et al., 2007), and energy saving behavior (e.g., Karlin et al., 2015). In particular, Kaiser et al. (2022) report that the average effect size of financial education interventions corresponds to about 0.123 standard deviation units. This is only about half as much as what we see for the impact of our media literacy intervention on the credibility of fakes, and roughly as much as what we observe for the impact on factual knowledge on average. However, Kaiser et al. (2022) stress that the effectiveness of education interventions diminishes over time. Given that the average effect size from their meta-analysis also includes interventions that were evaluated after several weeks or months, it is not surprising that the effect sizes from our media literacy intervention are comparably large.

A plausible explanation for the larger effectiveness of the media literacy relative to the fact checking intervention is that the former raises participants’ attention and enables them to critically evaluate the postings’ accuracy, whereas the latter only helps them to update beliefs about one specific fact and fails to enhance their more general skills. This matches the findings by Kaiser and Menkhoff (2022), who study different types of interventions offered to small-scale retailers in Uganda and show that “active learning” has a positive effect on savings and investment outcomes while traditional lecturing is ineffective. Comparable results on the advantages of active as opposed to passive learning have been documented in other domains, including science, technology, engineering, and maths (e.g., Deslauriers et al., 2011; Ruiz-Primo et al., 2011; Freeman et al., 2014). Relatedly, Drexler et al. (2014) show that a heuristics-based approach relying on “rules-of-thumb-training” – such as our ten tips to spot false news – generates larger behavioral impacts than the teaching of full curricula.

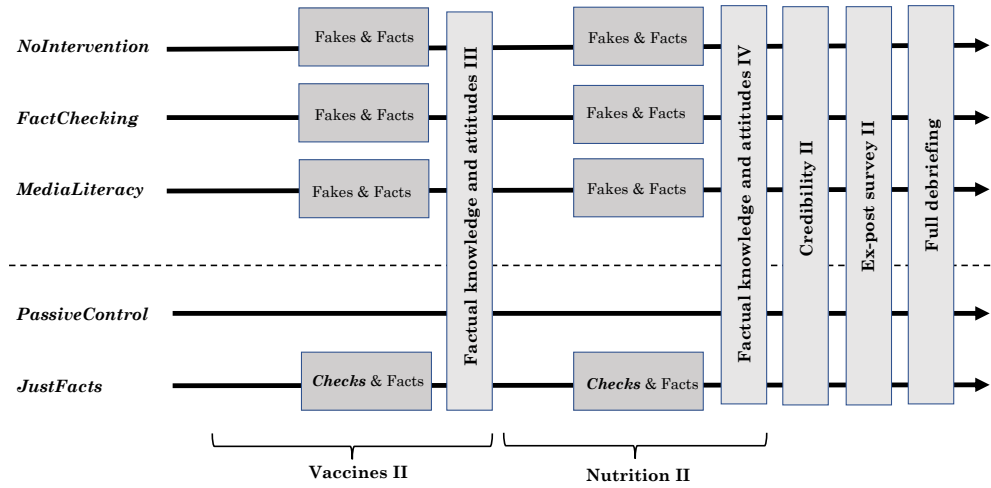
3. Experimental design

3.1. Survey flow

We start by randomizing the participants of our online survey experiment into one out of five groups of approximately equal size: (i) NOINTERVENTION, (ii) FACTCHECKING, (iii) MEDIALITERACY, (iv) JUSTFACTS, and (v) PASSIVECONTROL. To study both the short- and longer-term effects of our interventions, we conduct two waves of the experiment, where we re-invite the same participants one week after they completed Wave I and allocate them to the same treatment group as before. Figure 1 gives an overview of our survey flow, further details are discussed below.



(a) Overview Wave I



(b) Overview Wave II

Figure 1: Survey flow

3.1.1. Wave I

Baseline survey All participants start with a baseline survey on standard demographics like age, gender, family status, household income, education, profession, and personality traits (“big five”). In addition, we inquire participants’ prior knowledge on current events, health, and nutrition. To avoid priming effects on subsequent questions, we don’t include questions about the figures of interest after the interventions, but ask (i) how many days Joe Biden has been President of the United States, (ii) when to see a doctor in case of high temperature, and (iii) how many servings of fruit and vegetables are officially recommended per day. To measure the strength of participants’ prior beliefs, we also ask how certain they are about the accuracy of their responses on a 5-point Likert scale ranging from *Very uncertain* to *Very certain*.

Vaccines Next, participants in the NOINTERVENTION, the FACTCHECKING, and the MEDIALITERACY group are shown two pieces of “fake news” (“fakes” henceforth) and two facts on Corona vaccines in randomized order.⁴ The fakes and facts were manually collected from Facebook, i.e., we use screenshots of Facebook postings that actually exist. The collected fakes had to meet the following criteria: They were only included if we could find appropriate fact checks debunking the false information, they had to be either about COVID or nutrition and they had to be posted not too long ago. In addition, all fakes and facts had to contain a concrete numerical value (e.g., “50 people died after vaccination in a Sana clinic”) that we could later on ask for. Using this catalogue of criteria led to the pool of posts we use. See Appendix A.3 for all fakes and facts that are part of the experiment.

Participants in the NOINTERVENTION group do not receive further information. Participants in the FACTCHECKING group, in contrast, receive additional fact checks that explicitly debunk the false information (e.g., an official statement that the story about 50 deaths after vaccination in a Sana clinic is false). All fact checks stem from sources that are commonly perceived as trustworthy (e.g., *Correctiv.org*, a major German fact checking initiative). The fact checks are shown *prior* to the fakes that they correct. We thereby follow the current procedure on Facebook, where false or misleading information – if detected – is overlain with a warning message that redirects the user to a fact check; the original post can only be seen after the user closes the warning. Moreover, given that the timing of the intervention is known to be relevant (Lewandowsky et al., 2012; Brashier et al., 2021), displaying the fact check prior to the respective fake makes the fact checking better comparable to the media literacy intervention. See Appendix A.3 for all fact checks that we use.

Participants in the MEDIALITERACY group receive Facebook’s official “Tips to spot false news” *before* they are exposed to fakes and facts about Corona vaccines.⁵ These tips actually exist on the platform and comprise ten short pieces of advice, including “Be skeptical of catchy headlines”, “Look closely at the link”, and “Investigate the source”; Appendix A.1 shows the full list. We inform our participants that these tips have been developed by Facebook itself. We display one tip per page and ask the participants to read them carefully before they proceed to the Facebook postings on Corona vaccines.

⁴We find no evidence for order effects: Our main results are unaffected when we control for the order of the fakes and facts. Moreover, we find no differences between participants who first saw a fake and participants who first saw a fact.

⁵These tips have been developed in cooperation with several professional fact checking initiatives. See <https://www.facebook.com/help/188118808357379> (Dez 2021).

In contrast to the other groups, participants in the JUSTFACTS and in the PASSIVECONTROL group are *not* exposed to fakes. While the PASSIVECONTROL group does not see any postings at all, participants in the JUSTFACTS group receive the same two facts and fact checks (without the corresponding fakes) as participants in the FACTCHECKING group.⁶ We can thereby infer our participants’ average prior beliefs and attitudes from responses by the PASSIVECONTROL, and the impact of stand-alone fact checks from the JUSTFACTS group.

After exposure to the Facebook postings, we ask all participants four factual questions that are tailored to the fakes and facts just shown (e.g., we ask how many people died after vaccination in a Sana clinic).⁷ Each question asks for a specific number, and participants must give their answer through an input box, i.e., we do not provide a list of pre-defined options. To secure high quality answers, we use a bonus payment scheme that rewards participants whose answers are close to the true value.⁸

Next, we inquire all participants’ willingness to get vaccinated against Covid-19. We start by partitioning participants into those who are already fully vaccinated by the time of the experiment, and those who are not. Then, we ask the former group about their willingness to get a booster injection as soon as it is officially recommended, and the latter about their willingness to get vaccinated against Covid-19 in general. Answers could be given on a 5-point Likert scale ranging from *Very likely* to *Very unlikely*. To avoid experimenter demand effects, we do not incentivize this question with a potential bonus payment (see Section 6.5.3 for further discussion).

Nutrition The second part of Wave I is analogous to part one, except that we switch from Corona vaccines to nutritional topics, and that there are no further interventions (i.e., the setup is identical for participants in the NOINTERVENTION, the FACTCHECKING, and the MEDIALITERACY group). The main idea is to explore if the fact checking and the media literacy interventions stay effective in a different context that is health-related, too, but unlikely to be influenced by politically motivated reasoning.⁹ As before, participants in the NOINTERVENTION, the FACTCHECKING, and the MEDIALITERACY group are shown two fakes and two facts on nutritional topics in randomized order; these fakes and facts have to fulfill the same requirements as above. All fakes on nutrition promote the intake of needless dietary supplements such as extra protein or Vitamin C. Participants in the PASSIVECONTROL and the JUSTFACTS group are not exposed to fakes on nutrition, but the latter receive two facts and two fact checks.¹⁰

Analogous to part one of the survey, we proceed with a quiz that comprises four factual questions tailored to the fakes and facts that have just been shown. Again, each of those questions asks for a specific number, answers must be given through an input box, and we remind our participants of the potential bonus payment to incentivize high quality answers.

⁶The fact checks are self-contained and can stand on their own.

⁷The correct answer in this example is zero.

⁸More specifically, we use a quadratic scoring rule, whereby answers close to the true value increase participants’ chance to receive a bonus payment of 20 EUR.

⁹Though not as topical as Corona vaccines, the consumption of (needless) dietary supplements is an important concern. Recent surveys indicate that nearly 50% of all German adults have purchased dietary supplements within the last six months, but almost a third of them feels ill-informed about potential health risks that go along with their consumption (Verbraucherzentrale, 2022). Moreover, the consumption of dietary supplements does typically not go along with improved public health (Radimer et al., 2004) – quite the contrary – as dietary supplements are often either ineffective (DGE, 2012) or even harmful (Chiou et al., 2011).

¹⁰Note that the FACTCHECKING group does *not* receive these fact checks.

Finally, we inquire all participants’ willingness to consume dietary supplements, where they can respond on a 5-point Likert scale ranging from *Very likely* to *Very unlikely*.

Credibility Next, we inquire the perceived credibility of all fakes, facts, and fact checks. To this end, we display all postings again and let participants rate their credibility on a 5-point Likert scale ranging from *Very credible* to *Very incredible*. Participants are only asked about postings that they saw during the experiment, i.e., the FACTCHECKING group is asked about fakes, facts, and fact checks, the MEDIA LITERACY and the NO INTERVENTION groups are asked about fakes and facts, the JUSTFACTS group is asked about facts and fact checks, and the PASSIVE CONTROL group is not asked at all. We deliberately inquire the credibility of fakes, facts, and fact checks at this late stage of the experiment to avoid priming effects on the preceding questions. Moreover, to avoid experimenter demand effects, the credibility questions are not incentivized with a potential bonus payment, and we explicitly state that there is “no correct answer” and that we are “interested in [the participants’] personal opinion”.

Ex-post survey The ex-post survey serves to enhance our understanding of potential mechanisms as well as to collect information that we did not inquire earlier to avoid priming effects.

We proceed in three steps. First, we display all fakes, facts, and fact checks again and let participants indicate which elements of the postings they perceive as especially trustworthy or untrustworthy. We pre-define five types of elements to obtain comparable responses: (i) format and spelling, (ii) content as such, (iii) pictures, (iv) source and URL, and (v) verified account.¹¹ Participants can mark none, one, or several elements per posting and must complete a mandatory tutorial to get familiar with the procedure before they can proceed (see Appendix A.4 for an example). As with the credibility questions, participants are only asked about postings that they saw during the experiment.

Second, we ask further questions on participants’ political party preferences, social media usage, if they got vaccinated during the past ten years, and if they agree with the current Corona regulations. Moreover, we ask all participants if they searched for further information online.

Third, we conduct two list experiments à la Blair and Imai (2012) to check whether our main results on attitudes are driven by a “Bradley effect” (Hopkins, 2009), whereby participants conceal socially undesirable opinions and attitudes (see Section 6.5.3 for further details).

Debriefing At the end of Wave I, we debrief all participants by displaying the correct answers to the factual questions on Corona vaccines and nutrition.

3.1.2. Wave II

To study the effectiveness of our fact checking and media literacy interventions in the longer-run, we re-invite all participants after about one week to Wave II of the experiment. Wave II replicates the steps from Wave I, except that there are no further interventions (i.e., the setup is identical for the NO INTERVENTION, the FACT CHECKING, and the MEDIA LITERACY groups), no baseline survey, and that we use a different set of fakes, facts, and fact checks.

¹¹Note that some of the postings do not exhibit every element; e.g., some of them have no *verified account* label.

We conclude the survey with a full debriefing of all participants. To this end, we display the correct answers to all factual questions, show all fact checks to all fakes that were used during the survey, and provide links to trustworthy websites on Corona vaccines and nutrition, where the participants can get further information on these topics if they wish.¹²

3.2. Implementation

The experiment was programmed with the survey software *Qualtrics* and conducted in cooperation with *respondi*, a major commercial panel provider.¹³ We used e-mails to invite around 3,000 German participants to Wave I of the survey, i.e., around 600 participants per group.¹⁴ Participants had to be between 18 and 59 years old; conditional on that requirement, the sample is representative for the German population in terms of gender, age, and state of residency.¹⁵ Participants could use their smartphones, tablets, or desktop PCs to answer our questions. Those who completed the survey received the usual payment by *respondi* plus the potential bonus payment.

We conducted Wave I of the experiment between September 9th and September 29th in 2021, and Wave II between September 26th and October 27th. Participants received a re-invitation about one week after they completed Wave I. The minimum interval of actual participation in the two waves is equal to eight days, though, the median interval is equal to 15, and the mean interval equal to 17.6 days (see Section 6.5.2 for further discussion). The response rate in Wave II is equal to 83% – which is roughly equal to *respondi*’s average – and we find no evidence for differential attrition.¹⁶ At the time the experiment took place, all German adults had had the opportunity to get fully vaccinated (two injections), and policy makers and health experts were discussing whether and when a third injection would make sense.

3.3. Balance check

Table A.1 displays the means and standard deviations of all control variables for each treatment group. Since we use the NOINTERVENTION group as baseline in the subsequent analyses, we also conduct *t*-tests on the difference in means between the NOINTERVENTION and each of the other treatment groups, respectively.

We find that our sample is strongly balanced with respect to age, gender, family status, state of residence, consumption of dietary supplements, and prior knowledge on current events, health, and nutrition, but there are small differences between some of the treatment groups for household income, education, party preferences, and Corona vaccination status. To take these imbalances into account, we include the full set of pre-registered control variables into each of our regression

¹²Specifically, we suggest to visit the websites of the *Robert Koch Institut (RKI)* (URL: https://www.rki.de/DE/Home/homepage_node.html) and the *National Ministry of Health* (URL: <https://www.bundesgesundheitsministerium.de/>) further information on Corona vaccines, and to visit the website of the *German Agency for Nutrition* (URL: <https://www.dge.de/>) for further information on nutrition.

¹³Cooperating with professional panel providers such as *respondi* has become standard in economic research; see, e.g., Stantcheva (2021) and Alesina et al. (2022) for examples and <https://www.respondi.com/> for further details on *respondi*.

¹⁴The sample size is comparable to related studies, e.g. Barrera et al. (2020) and Henry et al. (2020).

¹⁵Although our participants are likely to encounter misinformation on Corona vaccines frequently in their every day lives, we exclude participants aged 60+ as an especially vulnerable group from our experiment.

¹⁶Response rates per group: NOINTERVENTION 82.36%, FACTCHECKING 84.51%, MEDIA LITERACY 83.36%, JUST-FACTS 78.90%, PASSIVECONTROL 84.62%.

analyses.¹⁷ Since we did not pre-register participants’ Corona vaccination status as a control, we only include it as a robustness check – with one exception (see Section 4.1.3) our results are unaffected.

3.4. Variables

Next, we aggregate our participants’ responses to the various fakes and facts and convert them into measures suitable for regression analyses. We also standardize responses to the prior knowledge questions and generate an indicator for participants’ uncertainty about them. Table A.2 provides summary statistics of all dependent variables that we use in the analysis.

Credibility We start by computing each participant’s mean response to the credibility questions on Corona vaccine fakes, Corona vaccine facts, nutrition fakes, and nutrition facts for each of the two waves, respectively (i.e., we compute eight mean responses per participant). Then, we define a dummy variable equal to one if the mean response indicates that the participant perceives the fakes or facts on average as *Credible* or *Very credible*.¹⁸ This aggregation level allows us to examine the treatment effect on fakes and facts separate from each other, whereby we can show that our interventions have no detrimental effect on facts.

Factual knowledge Next, we standardize participants’ responses to the factual knowledge questions. To this end, we first compute the absolute distance between each response and the correct answer. E.g., the correct answer to “How many people died after vaccination in a Sana clinic?” is equal to zero; if the participant’s response is “50”, the absolute distance between response and correct answer is equal to 50. To avoid distortion through outliers, we winsorize all distances to each question at their 95th percentile.¹⁹ Then, based on the entire sample, we standardize all winsorized distances to have a mean of zero and a standard deviation of one, which allows us to compare responses across questions. Finally, we aggregate participants’ responses by computing their mean standardized distance to the correct answer to the factual questions on Corona vaccine fakes, Corona vaccine facts, nutrition fakes, and nutrition facts for each of the two waves, respectively (i.e., we compute eight mean responses per participant again).

Attitudes To capture participants’ attitudes, we define a dummy that is equal to one if participant i states to be *Likely* or *Very likely* to get vaccinated or boosted against Covid-19 in each of the two waves, respectively. Analogously, we define a dummy equal to one if he or she states to be *Unlikely* or *Very unlikely* to consume dietary supplements in the near future.²⁰

Likes and dislikes To measure how much attention participants pay to the content of the fakes and facts and how critically they evaluate them, we count how many elements they marked as trustworthy (“like” for brevity) or untrustworthy (“dislike”) in each posting. Then, we compute

¹⁷The pre-registered control variables are age, gender, family status, state of residence, personality traits (“big 5”), household income, education, party preferences, and prior knowledge on current events, health, and nutrition.

¹⁸Our results are robust to alternative cutoffs.

¹⁹Note that we pre-registered our intention to winsorize each participant’s responses at their 95th percentile. We obtain similar results when we drop outliers beyond the 95th percentile, winsorize responses at their 99th percentile, or do not winsorize at all. The distribution is based on the entire sample.

²⁰Our results are robust to alternative cutoffs.

the absolute number of likes and dislikes for Corona vaccine fakes, Corona vaccine facts, nutrition fakes, and nutrition facts for each of the two waves, respectively.

Prior knowledge Analogous to factual knowledge, we make participants’ responses to each of the three prior knowledge questions better comparable by computing the standardized distance between a participant’s response and the correct answer. In addition, we compute an indicator that is equal to one if participant i is on average *Very uncertain*, *Uncertain*, or *Undecided* about his or her prior knowledge.

3.5. Analysis

In our baseline analysis, we use OLS to estimate the regression equation

$$y_{iw} = \beta_0 + \beta_1 TG_i + \beta_2 X_i + \varepsilon_{iw}, \quad (1)$$

where y_{iw} corresponds to an outcome of participant i in survey wave w as described above, TG_i denotes participant i ’s treatment group, and X_i is a vector of pre-registered control variables including age, gender, party preferences, religion, education, family status, household income, personality traits, state of residence, and prior knowledge on current events, health, and nutrition. The baseline category in TG_i is the NOINTERVENTION group, i.e., we compare participants who receive fakes and facts without further intervention to participants in each of the other treatment groups.

4. Results

The main purpose of our paper is to study whether and when fact checking and media literacy interventions are able to debunk fake news that circulate on social media. To this end, we focus on comparing the NOINTERVENTION to the FACTCHECKING and the MEDIALITERACY group here and defer supporting analyses of the JUSTFACTS and the PASSIVECONTROL group to Section 6.

We consider three types of outcomes: the credibility of fakes and facts, factual knowledge on the topics the fakes and facts are dealing with, and attitudes towards Corona vaccination and the intake of dietary supplements. The idea is to examine a coherent cognitive chain: Do the interventions reduce the credibility of fakes (but not of facts)? If yes, does that translate into better factual knowledge on the topics the fakes are dealing with? If yes, does that affect participants’ attitudes?

4.1. ITT analysis

We start by examining the Intention to Treat Effects (ITT) of our interventions (Section 6.5.4 presents an IV analysis). In particular, we demonstrate that the effectiveness of fact checking tends to be limited to the fakes that are corrected, while the media literacy intervention helps to distinguish between fakes and facts more generally. Figures A.8 to A.10 illustrate the results for each outcome, treatment, and wave of the survey; further details are presented below.

4.1.1. Credibility

Table 1 presents the regression results for the perceived credibility of fakes. Panel A shows the estimates from comparing the FACTCHECKING, and Panel B from comparing the MEDIALITERACY to the NOINTERVENTION group, respectively.

Our first main result is that the fact checking intervention reduces the credibility of fakes on Corona vaccines in Wave I of the survey. In Wave I, participants from the FACTCHECKING group are on average 7 to 8 percentage points less likely to perceive fakes on Corona vaccines as *Very credible* or *Credible* than participants from the NOINTERVENTION group. The estimate is statistically significant at the 1%-level, the effect size corresponds to about 15.5% of a standard deviation in the dependent variable in the baseline NOINTERVENTION group and to about 23.7% of its mean value. In contrast to that, we find no statistically significant differences between the FACTCHECKING and the NOINTERVENTION group for fakes that are not corrected by fact checks, i.e., fakes on Corona vaccines in Wave II of the survey and fakes on nutrition in either wave.

Second, we find that the media literacy intervention reduces the credibility of fakes more generally than fact checking. In Wave I, participants from the MEDIALITERACY group are about 10 percentage points less likely to consider fakes on Corona vaccines as *Very credible* or *Credible* than participants from the NOINTERVENTION group. The estimate is statistically significant at the 1%-level; the effect size corresponds to about 22.9% of a standard deviation in the dependent variable and to 35% of its baseline value, whereby the effect is even larger than for the FACTCHECKING group.²¹ Unlike fact checking, the media literacy intervention also reduces the credibility of all fakes on nutrition and of fakes on Corona vaccines in Wave II of the survey, although the latter effect is small and not statistically significant when we include our controls. The estimates for nutrition, however, correspond to about 20.4% of a standard deviation in the dependent variable in the baseline NOINTERVENTION group and to about 6.8% of its mean value in both waves of the survey.²²

Crucially, neither intervention reduces the credibility of facts (see Table A.3 in Appendix D), i.e., participants do not become more skeptical towards social media postings per se. Instead, our results indicate that the fact checking and, in particular, the media literacy intervention enhance participants' truth discernment (Pennycook and Rand, 2021), whereby they can better distinguish between false and correct information that they encounter online.

Truth discernment This result is further confirmed by an explicit analysis of truth discernment. In particular, we consider fakes and facts on Corona vaccines (nutrition) in Wave I (Wave II) of the survey in a single regression equation and estimate

$$cred_{ij} = \alpha_0 + \alpha_1 fake_j + \alpha_2 TG_i + \alpha_3 (TG_i * fake_j) + \theta X_i + e_{ij} \quad (2)$$

by OLS, where $cred_{ij}$ is a dummy equal to one if participant i perceived group j of news items (fakes or facts) on average as *Very credible* or *Credible*, TG_i is a treatment indicator with the NOINTERVENTION group as omitted category, and $fake_j$ indicates fakes (in contrast to facts).

²¹The difference between the FACTCHECKING and the MEDIALITERACY group is not statistically significant, though (two-sided t -test, $p = 0.176$).

²²Note that the perceived credibility of fakes on nutrition is much larger than for fakes on Corona vaccines. We further discuss this heterogeneity in Section 4.2.

The parameter α_1 thus corresponds to the average difference in the perceived credibility of fakes and facts – which we interpret as truth discernment – in the baseline NOINTERVENTION group. The parameter of interest is α_3 , which corresponds to the average difference in truth discernment between the NOINTERVENTION and the FACTCHECKING or the MEDIALITERACY group, respectively.

Table A.4 displays our results. Panel A shows that relative to facts, participants from the FACTCHECKING group are on average 6.8 percentage points less likely to perceive fakes on Corona vaccines as *Very credible* or *Credible* than participants from the NOINTERVENTION group. Hence, fact checking enhances participants’ truth discernment for the fakes that are explicitly targeted by our intervention. In contrast to that, we do not observe a statistically significant difference in truth discernment between the FACTCHECKING and the NOINTERVENTION group for Corona vaccines in Wave II of the survey or for nutrition in either wave.

Unlike fact checking, media literacy enhances participants’ truth discernment throughout the entire survey (Panel B). In particular, relative to facts, participants from the MEDIALITERACY group are on average between 6.6 and 9.8 percentage points less likely to perceive fakes on Corona vaccines as *Very credible* or *Credible* than participants from the NOINTERVENTION group. With the exception of Corona vaccines in Wave II of the survey, all estimates are highly statistically significant at the 1%-level.

4.1.2. Factual knowledge

Table 2 shows the regression results for participants’ factual knowledge on the topics the fakes are dealing with. Again, Panel A shows the estimates from comparing the FACTCHECKING, and Panel B from comparing the MEDIALITERACY to the NOINTERVENTION group, respectively.

Consistent with the results on credibility, Panel A shows that the fact checking intervention enhances participants’ factual knowledge on Corona vaccines in Wave I of the survey. Specifically, responses by the FACTCHECKING group are on average about 0.32 standard deviations closer to the correct answer than responses by the NOINTERVENTION group; the effect is statistically significant at the 1%-level.²³ Somewhat surprisingly, we also observe that participants from the FACTCHECKING group give better answers to the factual knowledge questions on nutrition in Wave I of the survey. According to our estimates, responses by the FACTCHECKING group are about 0.08 standard deviations closer to the correct answer than responses by the NOINTERVENTION group; the effect is statistically significant at the 5%-level. A potential explanation is that the presence of fact checking makes participants generally more cautious towards implausible information, although it does not affect the perceived credibility of fakes that are not explicitly corrected. This would be in line with recent findings by Barrera et al. (2020) and Nyhan et al. (2020), who show that fact checks can improve the accuracy of respondents’ factual beliefs, but fail to affect more deep-rooted perceptions and attitudes (see Section 4.1.3 for further discussion). Consistent with that, we find no statistically significant differences in factual knowledge on fakes between the FACTCHECKING and the NOINTERVENTION group in Wave II of the survey, where no further fact checks are shown.

²³Note that aggregating the standardized responses to Corona vaccine fakes, Corona vaccine facts, nutrition fakes, and nutrition facts in each wave of the survey causes the reported means and standard deviations in Table 2 to be unequal to zero and one, respectively.

Table 1: Credibility of fakes

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | -0.080 | -0.071 | -0.015 | -0.010 | -0.009 | 0.016 | -0.026 | -0.018 |
| | [0.025] | [0.024] | [0.018] | [0.017] | [0.031] | [0.030] | [0.024] | [0.024] |
| p-value | (0.001) | (0.004) | (0.385) | (0.553) | (0.769) | (0.582) | (0.282) | (0.460) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,225 | 1,225 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 |
| Panel B: Media literacy | | | | | | | | |
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | -0.104 | -0.105 | -0.060 | -0.061 | -0.052 | -0.042 | -0.072 | -0.068 |
| | [0.024] | [0.024] | [0.019] | [0.019] | [0.030] | [0.029] | [0.025] | [0.025] |
| p-value | (0.000) | (0.000) | (0.002) | (0.001) | (0.084) | (0.143) | (0.004) | (0.006) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,020 | 1,020 | 1,020 | 1,020 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 0.299 | 0.299 | 0.901 | 0.901 | 0.403 | 0.403 | 0.831 | 0.831 |
| Std.Dev. DV | 0.458 | 0.458 | 0.299 | 0.299 | 0.491 | 0.491 | 0.375 | 0.375 |

Notes: Table 1 shows the OLS estimates of comparing the NOINTERVENTION to the FACTCHECKING (Panel A) and to the MEDIALITERACY group (Panel B), respectively. The dependent variable is a dummy equal to one if participant i perceives the **fakes** on Corona vaccines and nutrition in Wave I and in Wave II of the survey on average as *Very credible* or *Credible*. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

Analogous to the results on credibility, Panel B shows that the media literacy intervention enhances participants’ factual knowledge more generally than fact checking. In Wave I of the survey, responses by the MEDIALITERACY group are about 0.22 standard deviations closer to the correct answer than responses by the NOINTERVENTION group for fakes on Corona vaccines, and about 0.08 standard deviations closer for fakes on nutrition. Both effects are statistically significant. The impact of media literacy is thus smaller than the impact of fact checking when the FACTCHECKING group receives correct information in addition to the fakes, and roughly equivalent when it does not.²⁴ However, unlike fact checking, the media literacy intervention could also improve participants’ factual knowledge on Corona vaccines in Wave II of the survey. Specifically, responses by the MEDIALITERACY group are about 0.14 standard deviations closer to the correct answer than responses by the NOINTERVENTION group; the effect is statistically significant at the 1%-level when we include our controls. In contrast to that, we find no statistically significant difference in factual knowledge on nutrition between the MEDIALITERACY and the NOINTERVENTION group in Wave II of the survey, although the estimates have the expected sign. One plausible explanation is that, according to Table 1, fakes on nutrition seem to be more credible on average than fakes on Corona vaccines. As a result, the tips to spot false news could be more difficult to apply, which in turn entails a smaller difference between the MEDIALITERACY and the NOINTERVENTION group. In addition, the impact of our intervention is likely to decay over time (e.g., Nyhan, 2021; Maertens et al., 2021), which further reduces the effect size in Wave II of the survey.

Similar to the results on credibility, Table A.5 in Appendix D shows that neither intervention reduces participants’ factual knowledge on topics that the facts are dealing with. Hence, both the fact checking and the media literacy intervention enhance participants’ factual knowledge on average. Moreover, Table A.6 shows that the fact checking intervention enhances participants’ truth discernment for Corona vaccines in Wave I of the survey, while there is no statistically significant difference in truth discernment between the FACTCHECKING and the NOINTERVENTION group otherwise. The media literacy intervention, in contrast, enhances truth discernment more generally. In particular, all estimates have the expected sign and almost all of them are statistically significant.

4.1.3. Attitudes

Table 3 displays the regression results for participants’ attitudes towards Corona vaccination and the intake of (needless) dietary supplements. Panel A displays the estimates from comparing the FACTCHECKING, and Panel B from comparing the MEDIALITERACY to the NOINTERVENTION group, respectively.

Panel A reveals that the impact of fact checking on participants’ attitudes is extremely limited. Although participants from the FACTCHECKING group are more likely to state that they are *Likely* or *Very likely* to get vaccinated or boosted against Covid-19 than participants from the NOINTERVENTION group, much of the effect is driven by the smaller proportion of fully vaccinated participants in the latter (see Section 3.3). In particular, we find that a participant’s decision to

²⁴The difference between the FACTCHECKING and the MEDIALITERACY group is statistically significant at the 5%-level for fakes on Corona vaccines in Wave I of the survey (two-sided t -test, $p = 0.034$) and weakly statistically significant at the 10%-level for fakes on nutrition in Wave II of the survey (two-sided t -test, $p = 0.097$).

Table 2: Distance to truth on topics covered by fakes

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | -0.322 | -0.319 | -0.080 | -0.080 | -0.048 | -0.051 | 0.021 | 0.030 |
| | [0.047] | [0.048] | [0.035] | [0.035] | [0.054] | [0.054] | [0.038] | [0.038] |
| p-value | (0.000) | (0.000) | (0.023) | (0.024) | (0.371) | (0.350) | (0.596) | (0.426) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,225 | 1,225 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 |
| Panel B: Media literacy | | | | | | | | |
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | -0.219 | -0.220 | -0.078 | -0.081 | -0.114 | -0.138 | -0.026 | -0.041 |
| | [0.048] | [0.048] | [0.037] | [0.036] | [0.053] | [0.052] | [0.038] | [0.037] |
| p-value | (0.000) | (0.000) | (0.032) | (0.027) | (0.031) | (0.009) | (0.488) | (0.264) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,020 | 1,020 | 1,020 | 1,020 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 0.299 | 0.299 | 0.901 | 0.901 | 0.403 | 0.403 | 0.831 | 0.831 |
| Std.Dev. DV | 0.458 | 0.458 | 0.299 | 0.299 | 0.491 | 0.491 | 0.375 | 0.375 |

Notes: Table 2 compares distance to truth on topics that the Corona vaccine and nutrition **fakes** are dealing with between participants from the FACTCHECKING (Panel A) and the MEDIALITERACY (Panel B) and the NOINTERVENTION group, respectively. All estimates are OLS estimates. The dependent variable is equal to participant i 's mean average standardized distance to the correct answer. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition.

get vaccinated in the past strongly predicts his or her intention to get vaccinated in the future. As a result, the estimated difference in the average willingness to get vaccinated between the FACTCHECKING and the NOINTERVENTION group shrinks and becomes statistically insignificant when we control for participants' Corona vaccination status in columns 3 and 8.²⁵ Similarly, we do not find any statistically significant difference in participants' willingness to consume dietary supplements in either wave of the survey. We must therefore conclude that the fact checking intervention – though effective in reducing the perceived credibility of and enhancing factual knowledge on the fakes that are being targeted – fails to affect participants' attitudes on average. This is consistent with earlier findings by Barrera et al. (2020), Swire et al. (2017), Nyhan et al. (2020), and Jerit and Zhao (2020), among others, who show that fact checking can help to create “a more informed citizenry” (Nyhan et al., 2020, p.942), but struggles to change more deep-rooted perceptions and attitudes such as which political party to support or, as in our context, whether to get vaccinated against Covid-19 or not.

In line with the results on credibility and factual knowledge, Panel B shows that the media literacy intervention is more effective in swaying participants' attitudes than fact checking. In particular, participants from the MEDIALITERACY group are 3.4 to 4.8 percentage points more likely to state that they are *Likely* or *Very likely* to get vaccinated or boosted against Covid-19 than participants from the NOINTERVENTION group. In contrast to fact checking, this difference remains statistically significant when we control for participants' Corona vaccination status.²⁶ The effect size corresponds to about 8.7% of a standard deviation in the dependent variable in the baseline NOINTERVENTION group and 4.2% of its mean value in Wave I of the survey, and to about 11.1% of a standard deviation in the dependent variable and 6.1% of its baseline value in Wave II. Roughly 85% of all participants report that they are willing to get vaccinated or boosted against Covid-19, though. The relatively small effect size could thus be explained by a small proportion of participants who can still be convinced to get the shot; Section 4.3 computes persuasion rates à la DellaVigna and Kaplan (2007) to further address this issue. The estimates for participants' willingness to consume needless dietary supplements are statistically insignificant. One potential explanation is that, unlike Corona vaccination, the intake of dietary supplements is typically based on year-long habits (Bailey et al., 2013), and even if media literacy could affect participants' attitudes, further effects from attitudes to habit change are typically modest (Verplanken and Orbell, 2022).

In sum, the results from Section 4.1 support the idea that the effect of fact checking tends to be limited to the fakes that are being corrected, while enhancing participants' media literacy helps them to distinguish between fakes and facts more generally. Hence, in an environment where not every message can be fact checked, media literacy interventions are likely to be more effective on average.

²⁵Note that the willingness to get vaccinated or boosted against Covid-19 is the only instance where the smaller proportion of fully vaccinated participants in the NOINTERVENTION group plays a role.

²⁶The differences between the the FACTCHECKING and the MEDIALITERACY group are not statistically significant, though.

Table 3: Attitudes towards Corona vaccination and the intake of dietary supplements

| Panel A: Fact checking | | | | | | | | | | |
|----------------------------------|--------------------|---------|---------|-------------|---------|--------------------|---------|---------|-------------|---------|
| | Wave I | | | | | Wave II | | | | |
| | Corona vaccination | | | Supplements | | Corona vaccination | | | Supplements | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| Fact checking | 0.047 | 0.038 | 0.017 | -0.001 | 0.003 | 0.057 | 0.048 | 0.034 | 0.011 | 0.021 |
| | [0.023] | [0.021] | [0.019] | [0.028] | [0.028] | [0.026] | [0.025] | [0.022] | [0.031] | [0.031] |
| p-value | (0.037) | (0.068) | (0.341) | (0.959) | (0.908) | (0.032) | (0.049) | (0.127) | (0.717) | (0.504) |
| Controls | no | yes | yes + | no | yes | no | yes | yes + | no | yes |
| <i>N</i> | 1,225 | 1,225 | 1,225 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 | 1,022 |
| Panel B: Media literacy | | | | | | | | | | |
| | Wave I | | | | | Wave II | | | | |
| | Corona vaccination | | | Supplements | | Corona vaccination | | | Supplements | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| Media literacy | 0.054 | 0.053 | 0.034 | -0.041 | -0.037 | 0.064 | 0.068 | 0.048 | -0.022 | -0.012 |
| | [0.022] | [0.021] | [0.019] | [0.028] | [0.028] | [0.026] | [0.025] | [0.022] | [0.031] | [0.031] |
| p-value | (0.016) | (0.012) | (0.077) | (0.148) | (0.188) | (0.015) | (0.006) | (0.033) | (0.486) | (0.696) |
| Controls | no | yes | yes + | no | yes | no | yes | yes + | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,231 | 1,020 | 1,020 | 1,020 | 1,020 | 1,020 |
| Baseline: No Intervention | | | | | | | | | | |
| Mean DV | 0.813 | 0.813 | 0.813 | 0.548 | 0.548 | 0.781 | 0.781 | 0.781 | 0.552 | 0.552 |
| Std.Dev. DV | 0.390 | 0.390 | 0.390 | 0.500 | 0.500 | 0.413 | 0.413 | 0.413 | 0.497 | 0.497 |

Notes: Table 3 presents the OLS estimates of comparing the NoINTERVENTION to the FACTCHECKING (Panel A) and to the MEDIA LITERACY group (Panel B), respectively. The dependent variable is a dummy equal to one if participant i states to be *Likely* or *Very likely* to get vaccinated or boosted against Covid-19, or *Unlikely* or *Very unlikely* to consume dietary supplements in the near future. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition. In columns 3 and 8 (“yes +”), we also control for participants’ Corona vaccination status.

4.1.4. Multiple hypothesis testing

Our ITT analysis examines two treatments and twelve different outcome variables; i.e., we consider twenty-four hypotheses in sum. To take potential over-rejection of null hypotheses into account, we conduct a Romano-Wolf multiple hypothesis correction (Romano and Wolf, 2005). This procedure uses resampling methods (e.g., bootstrap) to control for the so-called family-wise error rate, i.e., the probability of rejecting at least one true null hypothesis in the family of hypotheses under test. The Romano-Wolf procedure offers more power than the procedures by Bonferroni and Holm, and it is furthermore able to eliminate the so-called subset pivotality assumption from previous resampling-procedures such as Westfall-Young (Clarke et al., 2020).

Table 4 displays both the original and the Romano-Wolf p -values from our ITT analysis of fact checking (Panel A) and media literacy (Panel B), respectively. As expected, all p -values grow under the correction. Yet, our main results remain qualitatively intact in that the fact checking intervention is limited to the fake news that are targeted, while the media literacy intervention helps more generally. In particular, we observe that media literacy has a statistically significant impact on the credibility of both fakes on Covid-19 vaccines and nutrition in Wave I of the survey even after the Romano-Wolf correction, and that some of the effects in Wave II survive.

Table 4: Romano-Wolf Multiple Hypothesis Correction

| Panel A: Fact checking | | | | | | | | | | | | |
|-------------------------|--------|-------|---------|-----------|-------|--------|---------|-------|---------|-----------|-------|--------|
| | Wave I | | | | | | Wave II | | | | | |
| | Corona | | | Nutrition | | | Corona | | | Nutrition | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Suppl. | Cred. | Dist. | Vaccine | Cred. | Dist. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Original p -value | 0.004 | 0.000 | 0.341 | 0.553 | 0.024 | 0.909 | 0.582 | 0.350 | 0.127 | 0.460 | 0.426 | 0.504 |
| Romano-Wolf p -value | 0.070 | 0.000 | 0.973 | 0.973 | 0.304 | 0.973 | 0.973 | 0.973 | 0.802 | 0.973 | 0.973 | 0.973 |
| Panel B: Media literacy | | | | | | | | | | | | |
| | Wave I | | | | | | Wave II | | | | | |
| | Corona | | | Nutrition | | | Corona | | | Nutrition | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Suppl. | Cred. | Dist. | Vaccine | Cred. | Dist. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Original p -value | 0.000 | 0.000 | 0.077 | 0.001 | 0.027 | 0.188 | 0.143 | 0.009 | 0.033 | 0.006 | 0.264 | 0.700 |
| Romano-Wolf p -value | 0.001 | 0.000 | 0.642 | 0.028 | 0.323 | 0.878 | 0.816 | 0.134 | 0.354 | 0.097 | 0.948 | 0.973 |

Notes: Table 4 displays the original and the Romano-Wolf p -values from our ITT analysis. The Romano-Wolf p -values were obtained with the STATA package *rwolf2* (Clarke, 2021). Each specification includes the full set of control variables. In columns 3 and 9, we also include participants' vaccination status as additional control.

4.2. Heterogeneity in baseline beliefs

The average impact of our fact checking and media literacy interventions is likely to depend on participants' prior beliefs. In particular, if participants update their beliefs in a Bayesian fashion, those with prior beliefs that are most different from the presented information would change their beliefs the most (see, e.g., Fong et al., 2021, for empirical evidence). This section demonstrates that our interventions are indeed more effective for supporters of the AfD ("Alternative for Germany"), a far-right populist party known for spreading misinformation on Corona vaccines (e.g., Gensing, 2021). Specifically, we show that there is substantial effect heterogeneity between AfD

and non-AfD supporters for fakes on Corona vaccines, but not for fakes on nutrition, where participants' prior beliefs are much more alike. Given that the effect heterogeneity is limited to fakes on Corona vaccines, we focus on that topic here, and defer the results on nutrition to Tables A.7 and A.8 in Appendix D.²⁷

4.2.1. Fact checking

Table 5 displays the average impact of fact checking for AfD (Panel A) and non-AfD supporters (Panel B) on each of our main outcomes in each wave of the survey, respectively.

There are two main insights. First, each point estimate in Panel A is larger than its counterpart in Panel B, which means that the average impact of fact checking is stronger for AfD than for non-AfD supporters. With the exception of factual knowledge, each of these differences is highly statistically significant (two-sided t -tests, $p < 0.001$). Second, there is ample heterogeneity in the baselines, i.e., the mean dependent variables in the NOINTERVENTION group, which we interpret as participants' average prior beliefs: AfD supporters are on average almost twice as likely to perceive fakes on Corona vaccines as *Very credible* or *Credible* than non-AfD supporters, their responses to the factual knowledge questions are further away from the correct answer by roughly a third, and they are only half as likely to state that they are willing to get vaccinated or boosted against Covid-19. Hence, one explanation for the effect heterogeneity between AfD and non-AfD supporters is that the proportion of participants who can still update their beliefs is substantially larger for the former than for the latter group. We elaborate on this idea in Section 4.3, where we compute persuasion rates for each of our interventions. Differing trust in the fact checks, in contrast, does not seem to drive the effect heterogeneity: AfD supporters perceive the fact checks on Corona vaccines on average almost as credible as non-AfD supporters (see Section 6.4).

Column 3 shows that AfD supporters from the FACTCHECKING group are 13.7 percentage points more likely to state that they are willing to get vaccinated or boosted than AfD supporters from the NOINTERVENTION group. The effect is statistically significant at the 5%-level; it corresponds to 27.8% of a standard deviation in the dependent variable in the baseline NOINTERVENTION group and to 35.1% of its mean value. This result is especially remarkable given that fact checking typically fails to affect participants' attitudes (see Section 4.1.3). Here, the relatively strong impact of fact checking on the credibility of and factual knowledge on fakes, combined with the initially small average proportion of AfD supporters who want to get vaccinated or boosted (47.8% vs. 84.1% for the non-AfD supporters), is potent enough to sway attitudes of those AfD supporters who can still be convinced, while attitudes of non-AfD supporters remain unaffected.

4.2.2. Media literacy

Analogous to Table 5, Table 6 displays the average impact of media literacy for AfD (Panel A) and non-AfD supporters (Panel B) on each of our main outcomes in each wave of the survey, respectively. Again, the impact of our intervention is larger for AfD than for non-AfD supporters;

²⁷We do not find any systematic effect heterogeneity in terms of education, age, social media usage, support of policy measures to counteract the spread of the Corona virus, or prior knowledge on current events, health, and nutrition. When we compare participants who are fully vaccinated to those who are not, we find a similar, though less pronounced, pattern as for AfD supporters (see Section 6.5 for a detailed analysis).

Table 5: Heterogeneity in baseline beliefs on Corona vaccination – FACTCHECKING

| Panel A: Fact checking – AfD supporters | | | | | | |
|--|---------------|---------|---------|----------------|---------|---------|
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Vaccine |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Fact checking | -0.166 | -0.362 | 0.137 | -0.221 | -0.325 | 0.090 |
| | [0.099] | [0.171] | [0.062] | [0.096] | [0.178] | [0.081] |
| p-value | (0.095) | (0.036) | (0.029) | (0.024) | (0.072) | (0.267) |
| Controls | yes | yes | yes + | yes | yes | yes + |
| <i>N</i> | 114 | 115 | 115 | 99 | 99 | 99 |
| Baseline: No Intervention | | | | | | |
| Mean DV | 0.508 | 0.509 | 0.390 | 0.788 | 0.358 | 0.404 |
| Std.Dev. DV | 0.504 | 0.847 | 0.492 | 0.412 | 0.795 | 0.495 |
| Panel B: Fact checking – non-AfD supporters | | | | | | |
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Vaccine |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Fact checking | -0.055 | -0.316 | 0.006 | 0.046 | -0.032 | 0.030 |
| | [0.025] | [0.051] | [0.019] | [0.031] | [0.058] | [0.023] |
| p-value | (0.029) | (0.000) | (0.752) | (0.145) | (0.576) | (0.192) |
| Controls | yes | yes | yes + | yes | yes | yes + |
| <i>N</i> | 1,107 | 1,110 | 1,110 | 923 | 923 | 923 |
| Baseline: No Intervention | | | | | | |
| Mean DV | 0.277 | 0.386 | 0.825 | 0.369 | 0.331 | 0.779 |
| Std.Dev. DV | 0.448 | 0.840 | 0.381 | 0.483 | 0.850 | 0.415 |

Notes: Table 5 displays the effect heterogeneity between AfD supporters (Panel A) and non-AfD supporters (Panel B) for our **Fact checking** intervention. In columns 1 and 4, the dependent variable is a dummy equal to one if participant i perceives the fakes on Corona vaccines as *Very credible* or *Credible* on average. In columns 2 and 5, the dependent variable is equal to participant i 's mean average standardized distance to the correct answer. In columns 3 and 6, the dependent variable is a dummy equal to one if participant i states to be *Likely* or *Very likely* to get vaccinated or boosted against Covid-19. All estimates are OLS estimates. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition. In columns 3 and 6 ("yes +"), we also control for participants' Corona vaccination status.

with the exception of factual knowledge, each of these differences is highly statistically significant (two-sided t -tests, $p < 0.001$). We also observe large differences in participants’ baselines. Hence, part of the effect heterogeneity can be explained by the different proportion of participants who can still update their beliefs (see Section 4.3 for further discussion).

Similar to fact checking, we find that media literacy has a large positive impact on AfD supporters’ attitudes towards Corona vaccination. Specifically, AfD supporters from the MEDIALITERACY group are about 14.9 percentage points more likely to state that they are *Likely* or *Very likely* to get vaccinated or boosted against Covid-19 than AfD supporters from the NOINTERVENTION group. The effect is statistically significant at the 5%-level; it corresponds to about 30.3% of a standard deviation in the dependent variable in the baseline NOINTERVENTION group and to 38.2% of its mean value. In contrast to fact checking, this difference remains statistically significant in Wave II of the survey, and the effect size is comparable to Wave I. We thus conclude that media literacy is even more successful in swaying attitudes of participants who can still be convinced.

In sum, we find that both the fact checking and the media literacy intervention are more effective for AfD supporters, whose prior beliefs on Corona vaccines are on average further away from the truth and can thus be updated more strongly. This result stands in contrast to previous findings on motivated reasoning (e.g., Lewandowsky et al., 2012; Jerit and Zhao, 2020), whereby preexisting worldviews or attachments to a political party can impede efforts to debunk fake news. However, Pennycook et al. (2020, 2021) argue that it is often a lack of attention rather than partisanship that drives such results. Our evidence is consistent with the latter line of thought (see also Section 5), since both the fact checking and the media literacy intervention increase the awareness of fake news on Corona vaccines, and thereby induce participants to update their beliefs if they can still do so.

4.3. Persuasion rates

Sections 4.1 and 4.2 reveal that many participants are willing to get vaccinated or boosted against Covid-19 irrespective of the interventions. As a result, the ITT effects are relatively small. To adjust our estimates for the share of participants left to be convinced, we compute persuasion rates à la DellaVigna and Kaplan (2007), both for the full sample as well as for AfD and non-AfD supporters, respectively. To that end, we define the share of participants left to be convinced as the NOINTERVENTION group’s share of participants stating to be *Very unlikely*, *Unlikely*, or *Undecided* to get vaccinated or boosted on the 5-point Likert scale. Then, we divide the ITT estimates from our preferred specifications (“yes +”) in Tables 3, 5, and 6 by that share.

Table 7 shows that around 37.7% (39.2%) of our participants could still be persuaded in Wave I (Wave II) of the survey. Specifically, 74.6% of the AfD supporters and 33.8% of the non-AfD supporters could still be persuaded in Wave I, and 75.0% (35.2%) in Wave II. Hence, the fact checking intervention could convince about 4.5% (8.6%) of all persuadable participants in Wave I (Wave II) of the survey, whereas the media literacy intervention could convince 9% in Wave I and 12.2% in Wave II. These magnitudes are similar to those in comparable papers (e.g., Barrera et al., 2020, p.13). The persuasion rates for AfD supporters are considerably larger than the persuasion rates for non-AfD supporters. Hence, even if we account for differences in the share of participants left to be convinced, both interventions are more effective for AfD than for non-AfD

Table 6: Heterogeneity in baseline beliefs on Corona vaccination – MEDIA LITERACY

| Panel A: Media literacy – AfD supporters | | | | | | |
|---|---------------|---------|---------|----------------|---------|---------|
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Vaccine |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Media literacy | -0.134 | -0.279 | 0.149 | -0.161 | 0.090 | 0.136 |
| | [0.096] | [0.166] | [0.071] | [0.091] | [0.170] | [0.068] |
| p-value | (0.167) | (0.095) | (0.039) | (0.080) | (0.600) | (0.047) |
| Controls | yes | yes | yes + | yes | yes | yes + |
| <i>N</i> | 116 | 116 | 116 | 101 | 101 | 101 |
| Baseline | | | | | | |
| Mean DV | 0.508 | 0.509 | 0.390 | 0.788 | 0.358 | 0.404 |
| Std.Dev. DV | 0.504 | 0.847 | 0.492 | 0.412 | 0.795 | 0.495 |
| Panel B: Media literacy – non-AfD supporters | | | | | | |
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Vaccine |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Media literacy | -0.096 | -0.208 | 0.021 | -0.027 | -0.151 | 0.042 |
| | [0.024] | [0.050] | [0.020] | [0.031] | [0.056] | [0.023] |
| p-value | (0.000) | (0.000) | (0.280) | (0.379) | (0.007) | (0.075) |
| Controls | yes | yes | yes + | yes | yes | yes + |
| <i>N</i> | 1,115 | 1,115 | 1,115 | 919 | 919 | 919 |
| Baseline | | | | | | |
| Mean DV | 0.277 | 0.386 | 0.825 | 0.369 | 0.331 | 0.779 |
| Std.Dev. DV | 0.448 | 0.840 | 0.381 | 0.483 | 0.850 | 0.415 |

Notes: Table 6 displays the effect heterogeneity between AfD supporters (Panel A) and non-AfD supporters (Panel B) for our **Media literacy** intervention. In columns 1 and 4, the dependent variable is a dummy equal to one if participant *i* perceives the fakes on Corona vaccines as *Very credible* or *Credible* on average. In columns 2 and 5, the dependent variable is equal to participant *i*'s mean average standardized distance to the correct answer. In columns 3 and 6, the dependent variable is a dummy equal to one if participant *i* states to be *Likely* or *Very likely* to get vaccinated or boosted against Covid-19. All estimates are OLS estimates. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition. In columns 3 and 6 ("yes +"), we also control for participants' Corona vaccination status.

supporters.

A complementary explanation for our finding could be that it is *ceteris paribus* easier to convince the former. E.g., DellaVigna and Gentzkow (2010) argue that persuasion is more effective when receivers of novel information are less certain about the truth (p.654). Similarly, Kuklinski et al. (2000) distinguish between *misinformed* – those who have wrong beliefs and hold them firmly (p.792) – and *uninformed* citizens. Consistent with that, we find that AfD supporters are on average less certain about their prior knowledge on current events, health, and nutrition. In particular, when we regress the uncertainty indicator from Section 3.4 on the AfD dummy, the resulting estimate indicates that AfD supporters are on average 4.1 percentage points more likely to be uncertain about their prior knowledge than non-AfD supporters; the effect is statistically significant at the 5%-level.²⁸ Similarly, when we further divide the sample of AfD supporters into those who are uncertain about their prior knowledge and those who are not, we find that the ITT estimates are larger and more statistically significant for the former group. Hence, the effect heterogeneity between AfD and non-AfD supporters is likely to be driven both by wrong priors and by uncertainty about them. Our results thus show that it is of utmost importance not to give up on those with seemingly extreme opinions; rather, one should try to address exactly those people.

Table 7: Persuasion rates Corona vaccination

| Panel A: Fact checking | | | | | | |
|--------------------------------|---------------|-------|---------|----------------|-------|---------|
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Full | AfD | Non-AfD | Full | AfD | Non-AfD |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Persuasion rate | 0.045 | 0.184 | 0.018 | 0.086 | 0.120 | 0.085 |
| Share to be persuaded | 0.377 | 0.746 | 0.338 | 0.392 | 0.750 | 0.352 |
| <i>N</i> | 1,225 | 115 | 1,110 | 1,022 | 99 | 923 |
| Panel B: Media literacy | | | | | | |
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Full | AfD | Non-AfD | Full | AfD | Non-AfD |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Persuasion rate | 0.090 | 0.200 | 0.062 | 0.122 | 0.181 | 0.120 |
| Share to be persuaded | 0.377 | 0.746 | 0.338 | 0.392 | 0.750 | 0.352 |
| <i>N</i> | 1,231 | 116 | 1,115 | 1,020 | 101 | 919 |

Notes: Table 7 displays the persuasion rates for our fact checking (Panel A) and media literacy interventions (Panel B) for Wave I and Wave II of the survey, respectively. Columns 1 and 4 consider all participants in the NOINTERVENTION, FACTCHECKING, and MEDIALITERACY groups. Columns 2 and 5 consider only AfD supporters, columns 3 and 6 only non-AfD supporters. The *Share to be persuaded* corresponds to the proportion of participants in the NOINTERVENTION group stating to be *Very unlikely*, *Unlikely*, or *Undecided* to get vaccinated or boosted.

²⁸In contrast to that, there is no evidence that AfD supporters' prior knowledge on current events, health, and nutrition is worse than that of non-AfD supporters. The proportions of AfD and non-AfD supporters who are on the edge of being persuaded (i.e., stating to be *Undecided*) to get vaccinated or boosted are similar, too.

5. Mechanisms

Section 4.1 shows that the effectiveness of fact checking tends to be limited to the fakes that are being corrected, while media literacy helps to distinguish between fakes and facts more generally, both in the short- and in the longer-run. A reasonable explanation is that the media literacy intervention raises participants' attention towards the Facebook postings and enables them to critically evaluate the postings' accuracy. Fact checking, in contrast, turns participants into passive recipients of the specific corrections and thus fails to enhance their skills.

To support the plausibility of this mechanism, this section shows that participants from the MEDIALITERACY group are on average more likely to actively search for further information than participants from the NOINTERVENTION group; moreover, they become better in identifying untrustworthy elements in fakes and trustworthy elements in facts. Figures A.11 to A.13 illustrate our results; further details are given below.²⁹

5.1. Search for further information

We first examine if participants actively search for further information. To this end, we generate a dummy equal to one if participant i reports to have used the Internet to answer the factual knowledge questions on Corona vaccines and nutrition, and use this dummy as dependent variable in equation (1).

Consistent with the proposed mechanism, Table 8 shows that all estimates for the FACTCHECKING group are negative, but they are not statistically significant (Panel A). In contrast to that, all estimates for the MEDIALITERACY group are positive, and they are statistically significant at the 10%-level in Wave I of the survey (Panel B). Specifically, participants from the MEDIALITERACY group are 4.7 to 4.9 percentage points more likely to search for further information than participants from NOINTERVENTION group. The effect size corresponds to 9.8% of a standard deviation in the dependent variable in the baseline NOINTERVENTION group and to about 12.0% (10.2%) of its mean value for factual questions on Corona vaccines (nutrition).

5.2. Likes and dislikes

Next, we show that the media literacy intervention helps participants to identify untrustworthy elements in fakes and trustworthy elements in facts, whereby they can better distinguish between false and correct information that they encounter online. Therefore, we consider the absolute number of dislikes on fakes and likes on facts (see Section 3.4) and compare participants from the NOINTERVENTION to the FACTCHECKING and to the MEDIALITERACY group, respectively.

Table 9 confirms that the media literacy intervention induces participants to dislike fakes more often, while the fact checking intervention has no such effect. In particular, all estimates for fact checking are close to zero and statistically insignificant (Panel A), while the estimates for media literacy are positive and statistically significant at the 1%-level (Panel B). In Wave I of the survey, participants from the MEDIALITERACY group dislike on average 1.2 more elements in

²⁹We do not find any systematic differences between the NOINTERVENTION, the FACTCHECKING, and the MEDIALITERACY group in terms of the time spent with the fakes and facts. Hence, while the media literacy intervention enhances participants' ability to critically evaluate the Facebook postings, it does not induce them to engage with them for a longer period of time.

Table 8: Search for further information

| Panel A: Fact checking | | | | | | | | |
|-------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | -0.041 | -0.034 | -0.037 | -0.035 | -0.015 | -0.008 | -0.032 | -0.026 |
| | [0.028] | [0.027] | [0.025] | [0.025] | [0.030] | [0.030] | [0.028] | [0.028] |
| p-value | (0.144) | (0.208) | (0.150) | (0.165) | (0.628) | (0.777) | (0.261) | (0.338) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,225 | 1,225 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 |

| Panel B: Media literacy | | | | | | | | |
|--------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | 0.049 | 0.048 | 0.047 | 0.047 | 0.010 | 0.016 | 0.036 | 0.043 |
| | [0.028] | [0.028] | [0.027] | [0.026] | [0.030] | [0.030] | [0.029] | [0.029] |
| p-value | (0.083) | (0.085) | (0.083) | (0.079) | (0.733) | (0.586) | (0.215) | (0.136) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,020 | 1,020 | 1,020 | 1,020 |

| Baseline: No Intervention | | | | | | | | |
|----------------------------------|-------|-------|-------|-------|-------|-------|-------|-------|
| Mean DV | 0.400 | 0.400 | 0.307 | 0.307 | 0.361 | 0.361 | 0.295 | 0.295 |
| Std.Dev. DV | 0.490 | 0.490 | 0.462 | 0.462 | 0.481 | 0.481 | 0.456 | 0.456 |

Notes: Table 8 shows the OLS estimates of comparing the NOINTERVENTION to the FACTCHECKING (Panel A) and to the MEDIALITERACY group (Panel B), respectively. The dependent variable is a dummy equal to one if participant i reports to have used the Internet to respond to the factual knowledge questions (fakes and facts) on Corona vaccines and nutrition in Wave I and in Wave II, respectively. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

fakes on Corona vaccines and 0.7 more elements in fakes on nutrition than participants from the NOINTERVENTION group.³⁰ The effect size corresponds to 36.3% (38.5%) of a standard deviation in the dependent variable in the baseline NOINTERVENTION group for fakes on Corona vaccines (fakes on nutrition) and to 36.7% (52.5%) of its mean value. The intervention’s impact persists in Wave II of the survey. Specifically, participants from the MEDIALITERACY group dislike on average 0.9 more elements in fakes on Corona vaccines and 0.7 more elements in fakes on nutrition than participants from the NOINTERVENTION group; this corresponds to 21.7% (29.1%) of a standard deviation in the dependent variable for fakes on Corona vaccines (fakes on nutrition) and to 20.1% (41.2%) of its mean value.

Analogously, Table 10 confirms that the media literacy intervention induces participants to like facts more often, while the fact checking intervention has no such effect. In particular, all estimates for fact checking are close to zero and statistically insignificant (Panel A), while the estimates for media literacy are positive and statistically significant at the 1%-level (Panel B). In Wave I of the survey, participants from the MEDIALITERACY group like on average 1 more element in facts on Corona vaccines and 0.7 more elements in facts on nutrition than participants from the NOINTERVENTION group. These effects correspond to 33.5% (20.9%) of a standard deviation in the respective dependent variable in the baseline NOINTERVENTION group and to 38.3% (20.7%) of its mean value. Again, the impact of media literacy persists in Wave II of the survey. Specifically, participants from the MEDIALITERACY group like on average 0.8 more elements in facts on Corona vaccines and 0.7 more elements in facts on nutrition than participants from the NOINTERVENTION group; this corresponds to 38.5% (25.3%) of a standard deviation in the dependent variable for facts on Corona vaccines (facts on nutrition) and to 36.3% (22.8%) of its mean value.

Crucially, participants from the MEDIALITERACY group do not generally like and dislike more elements in the fakes and facts. In particular, the media literacy intervention does not increase the number of likes on fakes (Table A.9), and its impact on the number of dislikes on facts is small and limited to (arguably untrustworthy) emojis in facts on nutrition (Table A.10). Hence, consistent with the results from Section 4.1, the media literacy intervention does not make participants more skeptical towards social media postings per se, but rather helps them to distinguish between trustworthy and untrustworthy (elements of the) information.

6. Further analyses

This section provides further context for the interpretation of our main results. Specifically, we examine the PASSIVECONTROL group to study if our interventions can offset the damage that fake news are causing, we consider the JUSTFACTS group to further investigate the limited effectiveness of fact checking, we discuss potential heterogeneity in the fakes and fact checks that we show to our participants, and we provide a battery of robustness checks for our main results.

³⁰All differences between the FACTCHECKING and the MEDIALITERACY group are statistically significant at the 1%-level (two-sided t -tests, $p < 0.000$).

Table 9: Dislikes of fakes

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | -0.023 | -0.093 | -0.048 | -0.085 | 0.070 | -0.022 | 0.014 | -0.041 |
| | [0.190] | [0.184] | [0.099] | [0.097] | [0.264] | [0.256] | [0.138] | [0.135] |
| p-value | (0.902) | (0.610) | (0.627) | (0.383) | (0.792) | (0.931) | (0.917) | (0.759) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| N | 1,225 | 1,225 | 1,225 | 1,225 | 1,030 | 1,030 | 1,030 | 1,030 |
| Panel B: Media literacy | | | | | | | | |
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | 1.242 | 1.194 | 0.679 | 0.660 | 1.088 | 0.903 | 0.725 | 0.680 |
| | [0.210] | [0.197] | [0.109] | [0.106] | [0.290] | [0.271] | [0.163] | [0.155] |
| p-value | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.001) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| N | 1,231 | 1,231 | 1,231 | 1,231 | 1,026 | 1,026 | 1,026 | 1,026 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 3.249 | 3.249 | 1.256 | 1.256 | 4.469 | 4.469 | 1.651 | 1.651 |
| Std.Dev. DV | 3.293 | 3.293 | 1.714 | 1.714 | 4.144 | 4.144 | 2.337 | 2.337 |

Notes: Table 9 compares the absolute number of **dislikes on fakes** on Corona vaccines and nutrition for participants from the NOINTERVENTION to the FACTCHECKING (Panel A) and the MEDIALITERACY group (Panel B) in Wave I and Wave II of the survey, respectively. All estimates are OLS estimates. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

Table 10: Likes of facts

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | 0.075 | 0.026 | 0.011 | 0.009 | 0.042 | 0.023 | 0.077 | 0.051 |
| | [0.167] | [0.162] | [0.187] | [0.184] | [0.138] | [0.136] | [0.180] | [0.181] |
| p-value | (0.653) | (0.872) | (0.953) | (0.962) | (0.763) | (0.867) | (0.670) | (0.777) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| N | 1,225 | 1,225 | 1,225 | 1,225 | 1,030 | 1,030 | 1,030 | 1,030 |
| Panel B: Media literacy | | | | | | | | |
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | 1.059 | 0.998 | 0.741 | 0.672 | 0.873 | 0.809 | 0.758 | 0.686 |
| | [0.196] | [0.185] | [0.200] | [0.193] | [0.158] | [0.149] | [0.189] | [0.182] |
| p-value | (0.000) | (0.000) | (0.000) | (0.001) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| N | 1,231 | 1,231 | 1,231 | 1,231 | 1,026 | 1,026 | 1,026 | 1,026 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 2.607 | 2.607 | 3.239 | 3.239 | 2.231 | 2.231 | 3.012 | 3.012 |
| Std.Dev. DV | 2.975 | 2.975 | 3.205 | 3.205 | 2.103 | 2.103 | 2.711 | 2.711 |

Notes: Table 10 compares the absolute number of **likes on facts** on Corona vaccines and nutrition for participants from the NOINTERVENTION to the FACTCHECKING (Panel A) and the MEDIALITERACY group (Panel B) in Wave I and Wave II of the survey, respectively. All estimates are OLS estimates. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

6.1. Comparison to PassiveControl group

The main purpose of our paper is to study whether and to what extent fact checking and media literacy interventions are able to debunk fake news that circulate on social media. The relevant benchmark are thus participants from the NOINTERVENTION group, who are exposed to fakes and facts without further intervention. While this comparison illustrates the effectiveness of our interventions in an environment where all participants see fakes and facts, it does not uncover to what extent the interventions are able to reverse the harm the fakes are causing. To better interpret the magnitude of our main estimates in that regard, this section compares responses from the PASSIVECONTROL group – i.e., participants who did not see any fakes or facts from Facebook at all – to the FACTCHECKING and the MEDIALITERACY group, respectively. The smaller the difference between those groups, the more effective is the respective intervention in repealing the impact of fakes.³¹ We complement the analysis with a comparison of the PASSIVECONTROL to the NOINTERVENTION group, which provides a benchmark for the absolute impact of fakes. Figures A.9 and A.10 illustrate our results; further details are provided below.³²

6.1.1. Factual knowledge

Table A.11 reveals that exposure to fakes on Corona vaccines and nutrition substantially impairs participants’ factual knowledge, and that neither the fact checking nor the media literacy intervention can fully offset the effect. In particular, we find that (almost) all estimates are positive and statistically significant at the 1%-level, which means that responses by the FACTCHECKING, the MEDIALITERACY, and the NOINTERVENTION group are significantly further away from the correct answer than responses by the PASSIVECONTROL group.

There are two additional insights. First, in comparison to the absolute impact of fakes (Panel C), the average effectiveness of our interventions is relatively small. E.g., in Wave I of the survey, the fact checking intervention repeals less than 40% of the damage caused by fakes on Corona vaccines, and 80% of the damage caused by fakes on nutrition. Similarly, the media literacy intervention repeals just 26.2% of the damage caused by fakes on Corona vaccines, and 80% for nutrition. Thus, while both interventions improve factual knowledge relative to the NOINTERVENTION group – i.e., in an environment where all participants see fakes and facts – they do not reverse the harm of fakes entirely.

Second, the estimates for nutrition are smaller than the estimates for Corona vaccines, i.e., responses are on average more similar to the PASSIVECONTROL group. One explanation could be that the false information on nutrition is closer to participants’ average prior than the false information on Corona vaccines. This would also be in line with our finding that the perceived credibility of fakes on nutrition is larger than for fakes on Corona vaccines (see Section 4.1.1). An alternative (though less plausible) explanation could be that the fakes on Corona vaccines are more persuasive than the fakes on nutrition, whereby participants’ factual knowledge on the former is shifted further away from their prior than their factual knowledge on the latter.

³¹Our interventions do not overcompensate the impact of fakes.

³²Recall that participants in the PASSIVECONTROL group are not asked to rate the credibility of fakes and facts (see Section 3.1), since the idea of the PASSIVECONTROL group is to measure factual knowledge and attitudes when not being exposed to any of the fakes or facts. Therefore, we cannot compare their credibility outcomes measured in the main treatments with the PASSIVECONTROL group.

6.1.2. Attitudes

In contrast to factual knowledge, Table A.12 shows that exposure to fakes has a relatively small, if any, impact on participants' attitudes (Panel C), and that both the fact checking and, in particular, the media literacy intervention can effectively repeal that impact. Specifically, we find that (almost) all estimates are close to zero and statistically insignificant, which means that participants from the FACTCHECKING, the MEDIALITERACY, and the NOINTERVENTION group report a similar willingness to get vaccinated (or boosted) against Covid-19 as well as a similar willingness to consume (needless) dietary supplements as participants from the PASSIVECONTROL group. This confirms the results from Section 4, whereby the majority of participants wants to get vaccinated or boosted irrespective of the interventions, and whereby it is difficult to affect habit-based attitudes on dietary supplements. Hence, the modest impact of our interventions on participants' attitudes (see Table 3) can also be explained by the small absolute impact of fakes: if exposure to fakes does not sway participants' attitudes to begin with, there is nothing that the fact checking or the media literacy intervention could change.³³

6.2. Comparison to JustFacts group

Section 5 demonstrates that the media literacy intervention helps participants to better distinguish between fakes and facts, while fact checking fails to enhance their skills. A complementary explanation for the smaller effectiveness of fact checking is that the corrections often repeat false claims and thus induce “anchoring” (Tversky and Kahneman, 1974) or “continued influence effects” (Lewandowsky et al., 2012), whereby users' beliefs are biased towards the initially presented values. To study the role of such effects in our context, this Section compares factual knowledge of the JUSTFACTS group – i.e., participants who only saw facts and fact checks – to the NOINTERVENTION and the PASSIVECONTROL group, respectively. If anchoring effects do not play a role, participants from the JUSTFACTS group should have better factual knowledge than participants from the NOINTERVENTION and the PASSIVECONTROL group, because they are given the correct information. If anchoring effects exist, however, we expect the JUSTFACTS group's factual knowledge to be in between the NOINTERVENTION and the PASSIVECONTROL group. The closer their responses are to the former, the stronger are the anchoring effects. Figure A.9 illustrates our results, further details are discussed below.

Table A.13 reveals that participants from the JUSTFACTS have better factual knowledge on nutrition than participants from the PASSIVECONTROL, and better factual knowledge on both Corona vaccines and nutrition than participants from the NOINTERVENTION group.³⁴ All differences are statistically significant at the 1%- or at the 5%-level.

There are three potential explanations for these results that are not mutually exclusive. First, consistent with Tversky and Kahneman (1974) and Lewandowsky et al. (2012), repetition of the

³³This is in contrast to Barrera et al. (2020), who find that exposure to fake news is highly persuasive. However, Barrera et al. (2020) consider fake news on migration in France, while we consider fake news on Corona vaccines and nutrition. The diverging results could thus be driven by differences in the context of the fakes, i.e., it could be easier to sway participants' voting intentions than their attitudes on Corona vaccination and dietary supplements.

³⁴Due to a technical issue with one of the fact checks on nutrition in Wave I of the survey, we do not aggregate participants' responses to the factual knowledge questions but consider just the one functioning fact check instead.

false claims could stick in participants’ memory. In particular, while all fact checks on Corona vaccines restate the respective fake news, the fact checks on nutrition do not recast any false or misleading numbers. As a result, fact checking increases participants’ factual knowledge on nutrition (as measured by the PASSIVECONTROL group), but reduces factual knowledge on Corona vaccines (although the JUSTFACTS still performs significantly better than the NOINTERVENTION group).³⁵ Second, Section 6.4 reveals that participants from the JUSTFACTS group perceive the fact checks on nutrition as significantly more credible than the fact checks on Corona vaccines. Thus, it could be that the former exhibit a stronger impact on participants’ priors, whereby they update their beliefs more extensively. Third and relatedly, participants’ prior beliefs on nutrition could be less firm than their beliefs on Corona vaccines, and thereby more easy to sway. Similarly, as discussed in Section 6.1 above, their priors on nutrition could be worse than their priors on Corona vaccines (i.e., further away from the truth), leaving more room for improvement through the fact checks. In sum, our evidence is in line with “anchoring” or “continued influence effects” that constitute one potential drawback of fact checking, but we cannot exclude alternative explanations, either.

6.3. Heterogeneity of fakes

The effectiveness of our interventions is likely to depend on the specific fakes that we select for the experiment. In particular, the more credible the fakes, the harder it is to detect them, whereby the ITTs decrease. To explore this issue in more detail, Figure A.14 displays the (disaggregated) mean credibility of all fakes as given by the NOINTERVENTION group on a 5-point Likert Scale.

We find that there is substantial heterogeneity in the credibility of fakes. In particular, participants perceive the fakes on Corona vaccines on average as less credible than the fakes on nutrition. Moreover, there is heterogeneity within topics: the perceived credibility of fakes on Corona vaccines ranges from 1.65 to 2.28 and of fakes on nutrition from 2.75 to 3.41. This is roughly consistent with the results from Section 4; in particular, it could explain why the impact of our interventions tends to be larger for fakes on Corona vaccines than for fakes on nutrition.

6.4. Heterogeneity of fact checks

According to Jerit and Zhao (2020), trust in the authors of corrective messages is a crucial cause for their effectiveness. In our context, distrust in the (authors of the) fact checks could further explain why the fact checking is less effective than the media literacy intervention. To explore the plausibility of this explanation, Figure A.15 displays the (disaggregated) mean credibility of all fact checks on a 5-point Likert Scale.³⁶

We find that the mean credibility for fact checks on Corona vaccines in Wave I of the survey – i.e., those that were displayed to the FACTCHECKING group – is surprisingly low: the fact checks rate between 2.62 and 2.73 for participants of the FACTCHECKING, and between 2.81 and 2.88 for participants of the JUSTFACTS group. This is significantly less than for fact checks on Corona vaccines in Wave II of the survey (ratings between 3.15 and 3.27) and for fact checks on nutrition in either Wave (ratings between 3.60 and 3.79). One possible driver of these differences

³⁵This explanation would also be consistent with the salience effects documented by Barrera et al. (2020).

³⁶Recall that the FACTCHECKING group was only shown fact checks on Corona vaccines in Wave I, while the JUSTFACTS group saw fact checks on all topics in both waves of the survey.

could be heterogeneity in the source. E.g., while both fact checks on Corona vaccines in Wave II of the survey are released by *dpa*, Germany’s most renowned news wire, the fact checks on Corona vaccines in Wave I stem from *Correctiv* and *AFP*, respectively. Although these are generally considered as reliable fact checking organizations, they might be less known among the participants of our experiment, and thus perceived as less trustworthy. On the other hand, one of the fact checks on nutrition in Wave II of the survey comes “just” from an online platform and three from national public authorities, but Figure A.15 shows that there are just minor differences in their perceived credibility. Hence, while it seems plausible that small trust in fact checking contributes to its relative ineffectiveness, we cannot claim with certainty that the authors of the fact checks are crucial components in this.

6.5. Robustness checks

This section provides several analyses that support the robustness of our main results. In particular, we show that heterogeneity in terms of participants’ vaccination status and the time span between Waves I and II does not play a role, we present evidence from list experiments that support the validity of the participants’ self-reported attitudes, and we demonstrate that our results are robust to using an IV approach, where we take potential non-compliance into account.

6.5.1. Heterogeneity in terms of vaccination status

At the time of our experiment (September/October 2021), every German adult could be fully vaccinated against Covid-19 (two injections); yet, only about 81.8% of our participants reported to have taken the opportunity. This raises two potential concerns. First, participants who selected themselves into vaccination could be systematically different from those who did not, especially with respect to the questions on Corona vaccines. Second, we elicited participants’ attitudes towards Corona vaccination with two different questions – “willingness to get vaccinated” vs. “willingness to get boosted” on a 5-point Likert-scale, respectively (see Section 3)– and responses to those two questions might not be entirely comparable. Our main analyses address these concerns with robustness checks, where we add participants’ Corona vaccination status as an additional control (see Section 4). To further support our main findings, this section shows that they are robust to splitting the sample into participants who are fully vaccinated and those who are not.³⁷

Consistent with the robustness checks that we already conducted, Tables A.16 and A.17 show that the effect of our fact checking and media literacy interventions are similar for participants who are fully vaccinated and those who are not. The ITTs of fact checking (Table A.16), for instance, hardly differ between the two subsamples; the main difference is the reduced precision in Panel B, stemming from the small sample of non-vaccinated participants. Table A.17 reveals that the willingness to get vaccinated is larger for non-vaccinated than fully vaccinated participants in the MEDIALITERACY group. This result is comparable to our findings on effect heterogeneity for AfD and non-AfD supporters from Section 4.2 and could either stem from differences in the baseline (i.e., there are more non-vaccinated participants left to be convinced) or from the

³⁷For brevity, we only report the results for questions related to Corona vaccines. We do not find any systematic effect heterogeneity between participants who are fully vaccinated and those who are not with respect to questions on nutrition.

differently posed question. In sum, however, we conclude that heterogeneity between participants who are fully vaccinated and those who are not is at most a minor concern.

6.5.2. Time span between Waves I and II

As we describe in Section 3.2, all participants received a re-invitation to Wave II about one week after they completed Wave I of the survey. However, as they could re-start the survey at any time after that, the time span between actual participation in the two waves is quite heterogeneous and lies between 8 and 45 days.

We conduct two analyses to confirm that heterogeneity in the time span between Waves I and II is unlikely to affect our results. First, when we regress the number of days between participation in the two waves on our treatment indicators (with the NOINTERVENTION group as omitted category) plus the full set of controls, almost all estimates are close to zero and statistically insignificant.³⁸ Hence, our treatments do not influence when participants re-start the survey. Second, there is no evidence that differences in the time span between Waves I and II affect the magnitude of our main estimates. When we interact the number of days between participation in the two waves with the treatment indicators in equation (1), all interaction terms are close to zero and statistically insignificant. Similarly, when we split the sample at the median time span (= 15 days) and estimate equation (1) on the two subsamples, respectively, the resulting point estimates resemble those from Section 4.1.³⁹

6.5.3. List experiments refute experimenter demand bias

Participants' self-reported attitudes on Corona vaccines and dietary supplements might suffer from social desirability or experimenter demand bias if the participants anticipate that we as researchers are in favor of vaccination and against the consumption of needless dietary supplements. However, recent empirical evidence from both the behavioral economics as well as the political science literature suggests that experimenter demand biases are small, if they exist at all (De Quidt et al., 2018; Mummolo and Peterson, 2019). As argued in Section 3, we further minimize the risk of experimenter demand bias by *not* incentivizing the corresponding questions with a potential bonus payment. In addition to that, this section conducts two list experiments à la Blair and Imai (2012) – one for Corona vaccines, one for nutrition – to confirm that our participants do not conceal any socially undesirable opinions and attitudes.

For each list experiment, we randomly partition our participants into two groups. One group receives a list of five, the other group a list of six statements in random order, where the additional sixth statement is “I prefer not to get vaccinated against Covid-19.” (“I take dietary supplements.”) and the other five statements are about unrelated topics (see Appendix A.2 for the full lists). Then, we ask each participant how many of those statements he or she would agree with. Finally, we compute the difference in means between the two groups for the number of supported statements, which can be interpreted as the proportion of participants who indirectly

³⁸The estimate for the PASSIVECONTROL group is negative and weakly statistically significant at the 10%-level, indicating that participants re-started the survey about half a day earlier than participants from the NOINTERVENTION group.

³⁹Note that the time between actual participation in Wave I and II is endogenous, so all robustness checks from this section should be interpreted with caution.

concede that they do not want to get vaccinated (that they consume dietary supplements). If these proportions are substantially larger than the proportions that we directly elicit in the experiment, the self-reported attitudes might suffer from social desirability or experimenter demand bias.

Columns 1 to 4 in Table 11 show that the proportion of participants who directly report that they do not want to get vaccinated or boosted is similar (column 1) or even larger (column 3) than the proportion elicited through the list experiment. The proportions of participants who directly report to consume dietary supplements are smaller than the proportion that we elicit through the list experiment (columns 2 and 4), but the differences are small and could be driven by the slightly different questions in the main and in the list experiment (“How likely are you to consume dietary supplements in the near future?” vs. “I consume dietary supplements.”) In sum, there is no evidence for systematic experimenter demand bias with respect to the self-reported attitudes on Corona vaccination and dietary supplements when we consider the entire sample.⁴⁰

Next, we check for experimenter demand bias within our three main treatment groups. Given our experimental setup, experimenter demand bias is especially likely to occur for the FACTCHECKING group when we ask about their willingness to get vaccinated in Wave I of the survey, and for the MEDIALITERACY group throughout. However, as for the entire sample, the proportions of participants who directly state that they do not want to get vaccinated are even larger than the proportions of participants who indirectly say so both for the FACTCHECKING and for the MEDIALITERACY group, both in Wave I and in Wave II of the survey. Hence, we do not find evidence for experimenter demand bias here. We do observe that for the MEDIALITERACY group, the proportion of participants who just indirectly report to consume dietary supplements is larger than the proportion of participants who directly says so. However, this is also true for the FACTCHECKING and the NOINTERVENTION group, where experimenter demand bias is unlikely to occur, which suggests that the result is driven by unrelated issues. In sum, we can conclude that experimenter demand bias is unlikely to confound our main results.

6.5.4. IV analysis

Participants from the FACTCHECKING and the MEDIALITERACY group might skip the intervention by just quickly clicking through the survey. In this case, the ITT estimates would underestimate the average treatment effect. To take this into account, this section presents an IV approach, where we use participants’ time spent with the interventions to determine their actual treatment status and their random assignment to a treatment group as an instrument.

We proceed in two steps. First, we specify how much time it takes to properly engage with the interventions. To this end, we asked eleven Research Assistants to carefully read the two fact checks as well as the ten tips to spot false news and recorded how much time they need. We find that the minimum amount of time spent on the fact checking intervention is equal to 24.7, and

⁴⁰We further support the analysis with a sample split (see Table A.14 in Appendix D), where we consider participants who directly report that they are likely to get vaccinated (consume dietary supplements) on the one hand, and participants who report that they are unlikely to get vaccinated (consume dietary supplements) on the other. Reassuringly, we find that the proportion of participants who indirectly concede that they are not going to get vaccinated against Covid-19 is much larger for the latter than for the former group. Similarly, the proportion of participants who indirectly concede to consume dietary supplements is much larger for participants who directly report that they are likely to do so than for participants who report that they are not.

Table 11: List experiments

| | All participants | | | | No Intervention | | | | Fact Checking | | | | Media Literacy | | | |
|-----------------|------------------|--------|---------|--------|-----------------|--------|---------|--------|---------------|--------|---------|--------|----------------|--------|---------|--------|
| | Wave I | | Wave II | | Wave I | | Wave II | | Wave I | | Wave II | | Wave I | | Wave II | |
| | Vacc. | Suppl. | Vacc. | Suppl. | Vacc. | Suppl. | Vacc. | Suppl. | Vacc. | Suppl. | Vacc. | Suppl. | Vacc. | Suppl. | Vacc. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) | (13) | (14) | (15) | (16) |
| Direct question | 0.19 | 0.30 | 0.22 | 0.31 | 0.22 | 0.30 | 0.26 | 0.30 | 0.17 | 0.29 | 0.20 | 0.31 | 0.16 | 0.27 | 0.20 | 0.28 |
| List experiment | 0.16 | 0.38 | 0.08 | 0.33 | 0.21 | 0.34 | 0.11 | 0.31 | 0.14 | 0.37 | 0.07 | 0.45 | 0.06 | 0.40 | 0.01 | 0.26 |
| N | 3,051 | 3,051 | 2,525 | 2,525 | 618 | 618 | 509 | 509 | 607 | 607 | 513 | 513 | 613 | 613 | 511 | 511 |

Notes: In row 1, Table 11 displays the proportion of participants who in the main experiment **directly** report to be *Very unlikely* or *Unlikely* to get vaccinated or boosted against Covid-19 (columns 1 and 3) and who **directly** report to be *Very likely* or *Likely* to consume dietary supplements (columns 2 and 4). In row 2, Table 11 displays the respective indirectly elicited proportions from the list experiments.

the minimum amount of time spent on the media literacy intervention is equal to 38.9 seconds.

Second, we define D_i as a dummy variable that indicates participant i 's actual treatment status. In particular, D_i is equal to one if i spent at least 24.7 seconds with the fact checking or 38.9 seconds with the media literacy intervention. Thus, D_i is equal to zero for participants who did not spend a reasonable amount of time with their respective intervention, and for all participants in the NOINTERVENTION group.⁴¹ Equation (1) thus extends to

$$y_{iw} = \gamma_0 + \gamma_1 \widehat{D}_i + \gamma_2 X_i + \epsilon_i \quad (3)$$

$$D_i = \pi_0 + \pi_1 TG_i + \pi_2 X_i + u_i, \quad (4)$$

which we estimate by 2SLS.

We prefer using a binary (rather than a continuous) measure for participants' actual treatment status for two reasons. First, the impact of time spent with the interventions is likely to be discrete: participants need a certain minimum amount of time to understand and process the novel information, but any time spent beyond that is unlikely to yield further benefits. Second, it generally takes more time to engage with the media literacy than with the fact checking intervention. Hence, using a binary measure for participants' actual treatment status makes the regression results better comparable across treatment groups.

Table A.15 confirms that the 2SLS estimates of equations (3) and (4) are larger, but qualitatively similar to their counterparts from Section 4.1.⁴² Moreover, the coefficients for π_1 demonstrate that close to 70% of the FACTCHECKING, and 74% of the MEDIALITERACY group spent a considerable amount of time with their respective intervention. Skipping the interventions could be a larger concern outside the context of our experiment, though, especially if they disrupt users' consumption of social media. We further discuss this issue below.

7. Conclusion

We conduct a large-scale randomized survey experiment on the short- and longer-term effects of fact checking and media literacy interventions to demonstrate that the impact of fact checking tends to be limited to the fakes that are being corrected, whereas the media literacy intervention helps to distinguish between fakes and facts more generally, both in the short- and in the longer-run. A plausible mechanism for our result is that media literacy enables participants to critically evaluate social media postings, while fact checking turns them into passive recipients of the specific corrections and thus fails to enhance their skills. Hence, in an environment where not every claim can be fact checked, media literacy is likely to be more effective than fact checking on average.

Our paper promotes brief media literacy interventions as an effective tool to fight fake news and advances current policy debates along these lines. The European Union, for instance, has

⁴¹ Using the minimum amount of time spent with the interventions as our threshold is the most conservative choice. When we use the median or mean amount of time from the RA survey, the IV estimates become larger, but are qualitatively unaffected.

⁴² As further robustness checks for the MEDIALITERACY group, we replace D_i with a dummy that is (i) equal to one if participant i reports to have used the tips during the experiment and (ii) equal to one if participant i could recall the tips correctly. The 2SLS estimates are larger than their counterparts in Table A.15 in both cases, but qualitatively unaffected.

recently asserted media literacy as a pivotal tool to counter misinformation on social media⁴³, and the UNESCO has provided policy guidelines for digital media and information literacy⁴⁴. Our results strongly support such endeavors and suggest that official media literacy campaigns – which are relatively cheap, scalable, and easy-to-implement – could be a valuable complement to existing efforts like fact checking.

Media literacy as a means to fight fake news might outperform fact checking for three more reasons. First, many news items are not clearly fake or fact. While fact checkers can only debunk fake news that clearly provide wrong or misleading information, media literacy is likely to raise users’ skills and awareness to spot even more subtle types of misinformation. Second, the effectiveness of fact checking hinges on users’ trust in the fact checker. E.g., about 50% of Americans think that fact-checkers are biased (Allen et al., 2021). In times where users’ trust in renowned sources like public service broadcasters is crumbling, the provision of fact checks is likely to be futile, as those whom the intervention is supposed to target are often those who distrust the fact checker. Third and relatedly, with the rise of Artificial Intelligence (AI), users might have to increasingly rely on their own critical judgment of what they see rather than blindly trust the assessment of third parties. This applies even when AI is now also being used by researchers and journalists to *assist* the detection of fake news. Specifically, while AI is currently able to detect suspicious and check-worthy news items and match them to claims that have previously been fact checked, fact-checking as such is still based on manual assessment and far from being fully automated.⁴⁵

Our analysis has several limitations that open avenues for further research. First, the magnitude of our coefficients is likely to depend on the specific fakes, facts, and fact checks as well as on the topics that we selected for the experiment. E.g., some fakes are harder to detect than others, which is likely to reduce the effectiveness of our interventions. Similarly, users may be more or less well informed about different topics, whereby they are more or less likely to benefit from our interventions. Also, there exist different types of fact-checking (e.g., journalist-based vs. crowd-sourced fact-checking, see Allen et al. (2021)), whose effectiveness might vary in different contexts and depend on the recipient. Therefore, we consider the qualitative results as our most insightful findings and recommend to interpret the precise point estimates with caution.

Second, it is unclear how many – and especially which – users would actually engage with media literacy interventions. In particular, some users might perceive such trainings as a nuisance and consequently skip them. In addition, it could be that mostly users who are well informed anyway decide to take part in media literacy interventions, while users with poor priors – i.e., those for whom the training would be most effective – prefer to shirk them. Participation in media literacy interventions will ultimately depend on their design. However, even if such interventions fail to reach the entire population, it is worthwhile to enhance the skills even of a subset of users, and this should be preferred over not doing anything. Thus, while our paper stresses the high potential of media literacy interventions as a tool to fight fake news, the most efficient ways to implement such trainings must be examined in future research.

⁴³See <https://digital-strategy.ec.europa.eu/en/policies/media-literacy> (Aug 2022).

⁴⁴<https://www.unesco.org/en/communication-information/media-information-literacy/policy-strategy> (Aug 2022).

⁴⁵See EU Horizon, URL: <https://ec.europa.eu/research-and-innovation/en/horizon-magazine/can-artificial-intelligence-help-end-fake-news> (Aug 2023).

Third and relatedly, while we demonstrate that media literacy interventions could help users to better distinguish between false and correct information that they encounter online, we remain agnostic about the concrete implementation of such trainings on behalf of social media platforms. In particular, it is unclear if social media would be willing to set up regular interventions (e.g., in terms of pop-up windows that appear every few weeks) and what the ideal type of intervention would look like. The fact that Facebook has developed a set of “Tips to Spot False News” on its own behalf is encouraging, though, and suggests that social media might be willing to cooperate with academics and policy makers. The ideal type of intervention is likely to depend on the specific social media platform – e.g., users on TikTok may require different tips than users on Facebook – and promises to be an interesting field for future research.

References

- Alesina, Alberto, Armando Miano, and Stefanie Stantcheva (2022) “Immigration and Redistribution,” *The Review of Economic Studies*.
- Allcott, Hunt and Matthew Gentzkow (2017) “Social media and fake news in the 2016 election,” *Journal of Economic Perspectives*, **31** (2), 211–36.
- Allen, Jennifer, Antonio A Arechar, Gordon Pennycook, and David G Rand (2021) “Scaling up fact-checking using the wisdom of crowds,” *Science Advances*, **7** (36), eabf4393.
- Bailey, Regan L, Jaime J Gahche, Paige E Miller, Paul R Thomas, and Johanna T Dwyer (2013) “Why US adults use dietary supplements,” *JAMA Internal Medicine*, **173** (5), 355–361.
- Barrera, Oscar, Sergei Guriev, Emeric Henry, and Ekaterina Zhuravskaya (2020) “Facts, alternative facts, and fact checking in times of post-truth politics,” *Journal of Public Economics*, **182**, 104123.
- Blair, Graeme and Kosuke Imai (2012) “Statistical analysis of list experiments,” *Political Analysis*, **20** (1), 47–77.
- Bode, Leticia and Emily K Vraga (2015) “In related news, that was wrong: The correction of misinformation through related stories functionality in social media,” *Journal of Communication*, **65** (4), 619–638.
- Brashier, Nadia M, Gordon Pennycook, Adam J Berinsky, and David G Rand (2021) “Timing matters when correcting fake news,” *Proceedings of the National Academy of Sciences*, **118** (5), e2020043118.
- Bursztyn, Leonardo, Aakaash Rao, Christopher Roth, and David Yanagizawa-Drott (2023) “Opinions as facts,” *The Review of Economic Studies*, **90** (4), 1832–1864.
- Cheung, Alan CK and Robert E Slavin (2016) “How methodological features affect effect sizes in education,” *Educational Researcher*, **45** (5), 283–292.
- Chiou, Wen-Bin, Chao-Chin Yang, and Chin-Sheng Wan (2011) “Ironie effects of dietary supplementation: illusory invulnerability created by taking dietary supplements licenses health-risk behaviors,” *Psychological Science*, **22** (8), 1081–1086.
- Clarke, Damian (2021) “rwoolf2 implementation and flexible syntax,” *Accessible at <https://www.damianclarke.net/computation/rwoolf2.pdf>*.
- Clarke, Damian, Joseph P Romano, and Michael Wolf (2020) “The Romano–Wolf multiple-hypothesis correction in Stata,” *The Stata Journal*, **20** (4), 812–843.
- De Quidt, Jonathan, Johannes Haushofer, and Christopher Roth (2018) “Measuring and bounding experimenter demand,” *American Economic Review*, **108** (11), 3266–3302.
- DellaVigna, Stefano and Matthew Gentzkow (2010) “Persuasion: Empirical Evidence,” *Annual Review of Economics*, **2** (1), 643–669.

- DellaVigna, Stefano and Ethan Kaplan (2007) “The Fox News effect: Media bias and voting,” *The Quarterly Journal of Economics*, **122** (3), 1187–1234.
- Deslauriers, Louis, Ellen Schelew, and Carl Wieman (2011) “Improved learning in a large-enrollment physics class,” *Science*, **332** (6031), 862–864.
- Drexler, Alejandro, Greg Fischer, and Antoinette Schoar (2014) “Keeping it simple: Financial literacy and rules of thumb,” *American Economic Journal: Applied Economics*, **6** (2), 1–31.
- Fong, Jessica, Tong Guo, and Anita Rao (2021) “Debunking Misinformation about Consumer Products: Effects on Beliefs and Purchase Behavior,” *SSRN*.
- Freeman, Scott, Sarah L Eddy, Miles McDonough, Michelle K Smith, Nnadozie Okoroafor, Hannah Jordt, and Mary Pat Wenderoth (2014) “Active learning increases student performance,” *Proceedings of the National Academy of Sciences*, **111** (23), 8410–8415.
- Sherridan, B.J. and Smith, B. (2020). How often does active learning actually occur? Perception versus reality, *AEA Papers and Proceedings*, 110: 304–308e in science, engineering, and mathematics,” *Proceedings of the National Academy of Sciences*, **111** (23), 8410–8415.
- Fryer Jr, Roland G (2017) “The production of human capital in developed countries: Evidence from 196 randomized field experiments,” in *Handbook of economic field experiments*, **2**: Elsevier, 95–322.
- Gensing, Patrick (2021) “Wie die AfD Angst vor Impfungen schürt,” *Tagesschau.de*.
- Grinberg, Nir, Kenneth Joseph, Lisa Friedland, Briony Swire-Thompson, and David Lazer (2019) “Fake news on Twitter during the 2016 US presidential election,” *Science*, **363** (6425), 374–378.
- Guess, Andrew M, Michael Lerner, Benjamin Lyons, Jacob M Montgomery, Brendan Nyhan, Jason Reifler, and Neelanjan Sircar (2020) “A digital media literacy intervention increases discernment between mainstream and false news in the United States and India,” *Proceedings of the National Academy of Sciences*, **117** (27), 15536–15545.
- Guess, Andrew, Jonathan Nagler, and Joshua Tucker (2019) “Less than you think: Prevalence and predictors of fake news dissemination on Facebook,” *Science Advances*, **5** (1), eaau4586.
- Guess, Andrew, Brendan Nyhan, and Jason Reifler (2018) “Selective exposure to misinformation: Evidence from the consumption of fake news during the 2016 US presidential campaign,” *European Research Council*, **9** (3), 4.
- Henry, Emeric, Ekaterina V Zhuravskaya, SM Guriev et al. (2020) *Checking and sharing alt-facts*: Centre for Economic Policy Research.
- Hill, Carolyn J, Howard S Bloom, Alison Rebeck Black, and Mark W Lipsey (2008) “Empirical benchmarks for interpreting effect sizes in research,” *Child development perspectives*, **2** (3), 172–177.
- Hopkins, Daniel J (2009) “No more Wilder effect, never a Whitman effect: When and why polls mislead about black and female candidates,” *The Journal of Politics*, **71** (3), 769–781.

- Jerit, Jennifer and Yangzi Zhao (2020) “Political misinformation,” *Annual Review of Political Science*, **23**, 77–94.
- Kaiser, Tim, Annamaria Lusardi, Lukas Menkhoff, and Carly Urban (2022) “Financial education affects financial knowledge and downstream behaviors,” *Journal of Financial Economics*, **145** (2), 255–272.
- Kaiser, Tim and Lukas Menkhoff (2022) “Active learning improves financial education: Experimental evidence from Uganda,” *Journal of Development Economics*, **157**, 102870.
- Karlin, Beth, Joanne F Zinger, and Rebecca Ford (2015) “The effects of feedback on energy conservation: A meta-analysis,” *Psychological bulletin*, **141** (6), 1205.
- Kuklinski, James H, Paul J Quirk, Jennifer Jerit, David Schwieder, and Robert F Rich (2000) “Misinformation and the currency of democratic citizenship,” *The Journal of Politics*, **62** (3), 790–816.
- Lazer, David MJ, Matthew A Baum, Yochai Benkler, Adam J Berinsky, Kelly M Greenhill, Filippo Menczer, Miriam J Metzger, Brendan Nyhan, Gordon Pennycook, and David Rothschild (2018) “The science of fake news,” *Science*, **359** (6380), 1094–1096.
- Lewandowsky, Stephan, Ullrich KH Ecker, Colleen M Seifert, Norbert Schwarz, and John Cook (2012) “Misinformation and its correction: Continued influence and successful debiasing,” *Psychological science in the public interest*, **13** (3), 106–131.
- Loewenstein, George and Zachary Wojtowicz (2023) “The Economics of Attention,” *Available at SSRN 4368304*.
- Luca, Michael (2015) “User-generated content and social media,” in *Handbook of Media Economics*, **1**: Elsevier, 563–592.
- Maertens, Rakoen, Jon Roozenbeek, Melisa Basol, and Sander van der Linden (2021) “Long-term effectiveness of inoculation against misinformation: Three longitudinal experiments,” *Journal of Experimental Psychology: Applied*, **27** (1), 1.
- Mummolo, Jonathan and Erik Peterson (2019) “Demand effects in survey experiments: An empirical assessment,” *American Political Science Review*, **113** (2), 517–529.
- Noar, Seth M, Christina N Benac, and Melissa S Harris (2007) “Does tailoring matter? Meta-analytic review of tailored print health behavior change interventions,” *Psychological bulletin*, **133** (4), 673.
- Nyhan, Brendan (2021) “Why the backfire effect does not explain the durability of political misperceptions,” *Proceedings of the National Academy of Sciences*, **118** (15), e1912440117.
- Nyhan, Brendan, Ethan Porter, Jason Reifler, and Thomas J Wood (2020) “Taking fact-checks literally but not seriously? The effects of journalistic fact-checking on factual beliefs and candidate favorability,” *Political Behavior*, **42** (3), 939–960.

- Nyhan, Brendan and Jason Reifler (2010) “When corrections fail: The persistence of political misperceptions,” *Political Behavior*, **32** (2), 303–330.
- (2015) “Does correcting myths about the flu vaccine work? An experimental evaluation of the effects of corrective information,” *Vaccine*, **33** (3), 459–464.
- Pennycook, Gordon, Ziv Epstein, Mohsen Mosleh, Antonio A Arechar, Dean Eckles, and David G Rand (2021) “Shifting attention to accuracy can reduce misinformation online,” *Nature*, **592** (7855), 590–595.
- Pennycook, Gordon, Jonathon McPhetres, Yunhao Zhang, Jackson G Lu, and David G Rand (2020) “Fighting COVID-19 misinformation on social media: Experimental evidence for a scalable accuracy-nudge intervention,” *Psychological Science*, **31** (7), 770–780.
- Pennycook, Gordon and David G Rand (2019) “Lazy, not biased: Susceptibility to partisan fake news is better explained by lack of reasoning than by motivated reasoning,” *Cognition*, **188**, 39–50.
- (2021) “The Psychology of Fake News,” *Trends in Cognitive Sciences*, **25** (5), 388–402.
- Radimer, Kathy, Bernadette Bindewald, Jeffery Hughes, Bethene Ervin, Christine Swanson, and Mary Frances Picciano (2004) “Dietary supplement use by US adults: data from the National Health and Nutrition Examination Survey, 1999–2000,” *American Journal of Epidemiology*, **160** (4), 339–349.
- Romano, Joseph P and Michael Wolf (2005) “Stepwise multiple testing as formalized data snooping,” *Econometrica*, **73** (4), 1237–1282.
- Rooney, Brenda L and David M Murray (1996) “A meta-analysis of smoking prevention programs after adjustment for errors in the unit of analysis,” *Health education quarterly*, **23** (1), 48–64.
- Roozenbeek, Jon, Sander van der Linden, Beth Goldberg, Steve Rathje, and Stephan Lewandowsky (2022) “Psychological inoculation improves resilience against misinformation on social media,” *Science Advances*, **8** (34), eabo6254, URL: <https://www.science.org/doi/abs/10.1126/sciadv.abo6254>.
- Ruiz-Primo, Maria Araceli, Derek Briggs, Heidi Iverson, Robert Talbot, and Lorrie A Shepard (2011) “Impact of undergraduate science course innovations on learning,” *Science*, **331** (6022), 1269–1270.
- Stantcheva, Stefanie (2021) “Understanding tax policy: How do people reason?” *The Quarterly Journal of Economics*, **136** (4), 2309–2369.
- Swire, Briony, Adam J Berinsky, Stephan Lewandowsky, and Ullrich KH Ecker (2017) “Processing political misinformation: comprehending the Trump phenomenon,” *Royal Society Open Science*, **4** (3), 160802.
- Tversky, Amos and Daniel Kahneman (1974) “Judgment under Uncertainty: Heuristics and Biases: Biases in judgments reveal some heuristics of thinking under uncertainty,” *Science*, **185** (4157), 1124–1131.

- Verplanken, Bas and Sheina Orbell (2022) “Attitudes, habits, and behavior change,” *Annual Review of Psychology*.
- Vosoughi, Soroush, Deb Roy, and Sinan Aral (2018) “The spread of true and false news online,” *Science*, **359** (6380), 1146–1151.
- Vraga, Emily K and Leticia Bode (2017) “Using expert sources to correct health misinformation in social media,” *Science Communication*, **39** (5), 621–645.
- Zhuravskaya, Ekaterina, Maria Petrova, and Ruben Enikolopov (2020) “Political effects of the internet and social media,” *Annual Review of Economics*, **12**, 415–438.

A. Supplementary material

A.1. Ten tips to spot false news

1. **Be skeptical of headlines.** False news stories often have catchy headlines in all caps with exclamation points. If shocking claims in the headline sound unbelievable, they probably are.
2. **Look closely at the link.** A phony or look-alike link may be a warning sign of false news. Many false news sites mimic authentic news sources by making small changes to the link. You can go to the site to compare the link to established sources.
3. **Investigate the source.** Ensure that the story is written by a source that you trust with a reputation for accuracy. If the story comes from an unfamiliar organization, check their "About" section to learn more.
4. **Watch for unusual formatting.** Many false news sites have misspellings or awkward layouts. Read carefully if you see these signs.
5. **Consider the photos.** False news stories often contain manipulated images or videos. Sometimes the photo may be authentic, but taken out of context. You can search for the photo or image to verify where it came from.
6. **Inspect the dates.** False news stories may contain timelines that make no sense, or event dates that have been altered.
7. **Check the evidence.** Check the author's sources to confirm that they are accurate. Lack of evidence or reliance on unnamed experts may indicate a false news story.
8. **Look at other reports.** If no other news source is reporting the same story, it may indicate that the story is false. If the story is reported by multiple sources you trust, it's more likely to be true.
9. **Is the story a joke?** Sometimes false news stories can be hard to distinguish from humor or satire. Check whether the source is known for parody, and whether the story's details and tone suggest it may be just for fun.
10. **Some stories are intentionally false.** Think critically about the stories you read, and only share news that you know to be credible.

A.2. List experiments

This section displays translations of the statements that we used in our list experiments on Corona vaccination and dietary supplements in Wave I and Wave II of our survey, respectively. The statements were shown in randomized order. The statements in regular font were shown to every participant, the statements in bold font only to about 50% of them. Assignment to see the additional statement was random.

List experiment on **Corona vaccination, Wave I:**

- I do not eat meat.
- I like football.
- I listen to the news on the radio in the morning.
- I live in a relatively small town.
- I usually go to bed rather late.
- **I prefer not to get vaccinated against Covid-19.**

List experiment on **dietary supplements, Wave I:**

- I like to go dancing.
- I work part-time but would prefer to work more.
- I like winter time.
- I do not have any pets.
- I suffer from a pollen allergy.
- **I consume dietary supplements.**

List experiment on **Corona vaccination, Wave II:**

- I like to go for a walk.
- I drink a lot of coffee.
- I do not have any siblings.
- My apartment is on the first floor.
- I like to eat bananas.
- **I prefer not to get vaccinated against Covid-19.**

List experiment on **dietary supplements, Wave II:**

- I went to university.
- I like to travel to Croatia.
- I do not have any kids.
- I like reading and I read a lot.
- I have a driving license.
- **I consume dietary supplements.**

A.3. Fakes, facts, and fact checks

Thema Sana-Kliniken



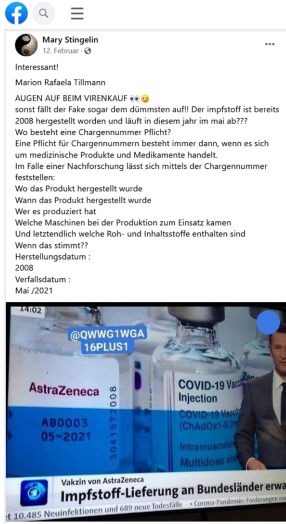
(a) Fake on Corona vaccines, Wave I

Thema Radioaktivität



(b) Fake on Corona vaccines, Wave I

Thema Impfpumpe



(c) Fake on Corona vaccines, Wave II

Thema Impfstoffhersteller



(d) Fake on Corona vaccines, Wave II

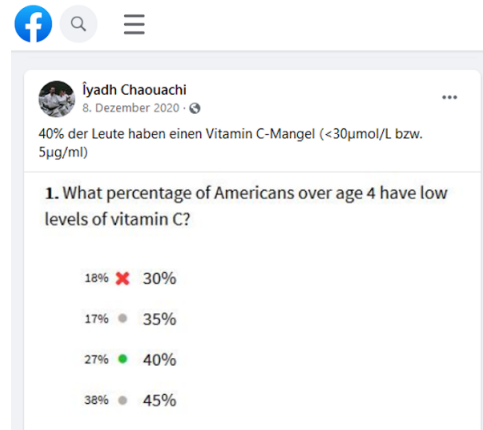
Figure A.1: Fakes on Corona vaccines.

Thema Proteinbedarf



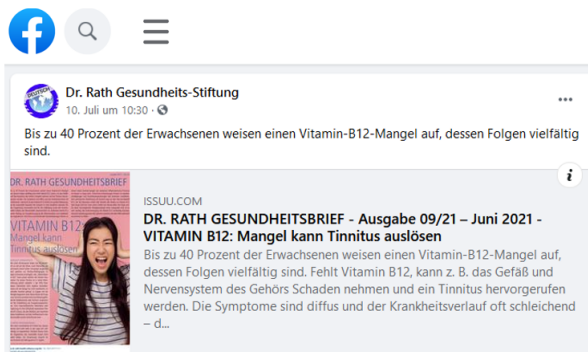
(a) Fake on Nutrition, Wave I

Thema Vitamin C



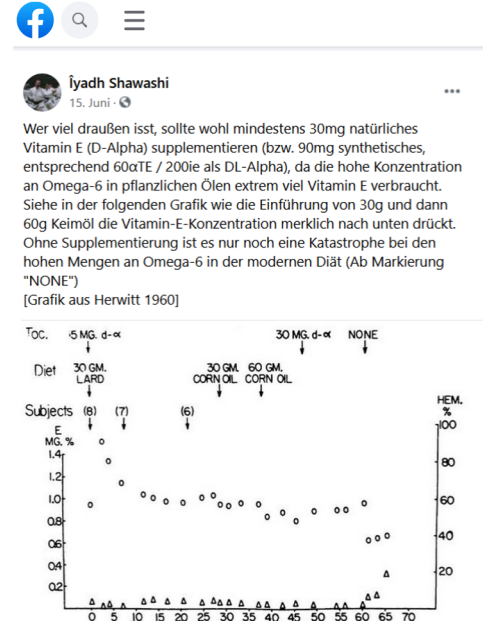
(b) Fake on Nutrition, Wave I

Thema Vitamin B12



(c) Fake on Nutrition, Wave II

Thema Vitamin E



(d) Fake on Nutrition, Wave II

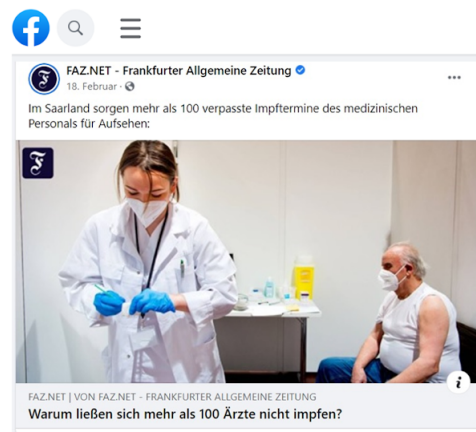
Figure A.2: Fakes on nutrition.

Thema Sputnik V



(a) Fact on Corona vaccines, Wave I

Thema Fernbleiben



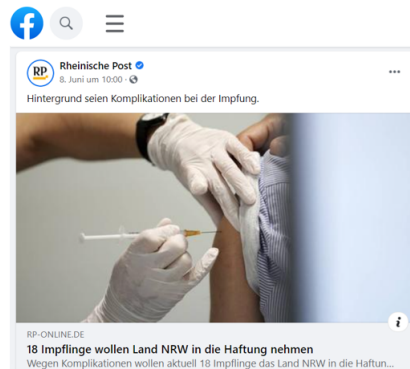
(c) Fact on Corona vaccines, Wave II

Thema Wirksamkeit



(b) Fact on Corona vaccines, Wave I

Thema Impfkomplicationen



(d) Fact on Corona vaccines, Wave II

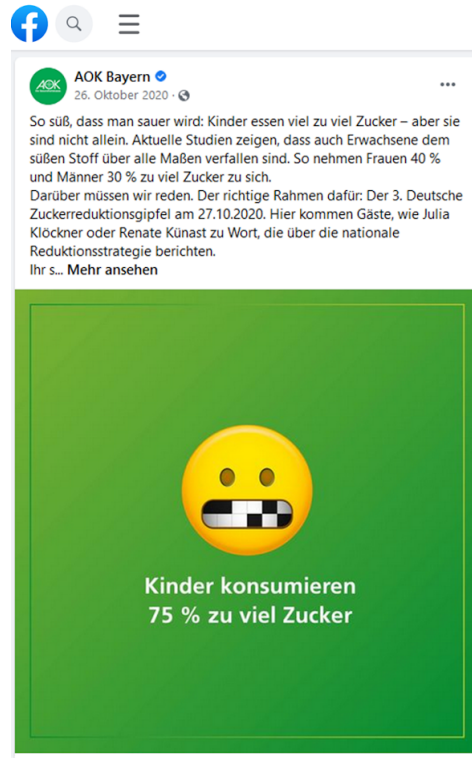
Figure A.3: Facts on Corona vaccines.

Thema Plastik



(a) Fact on Nutrition, Wave I

Thema Zucker



(b) Fact on Nutrition, Wave I

Thema Bewegung



(c) Fact on Nutrition, Wave II

Thema Nährwerte



(d) Fact on Nutrition, Wave II

Figure A.4: Facts on Nutrition.

Thema Sana-Kliniken

Fakten-Check



Gerücht über Todesfälle nach Corona-Impfungen im Sana-Klinikum ist frei erfunden

Vor allem auf Facebook und Telegram kursiert aktuell das Gerücht, dass 50 Mitarbeitende im Berliner Sana-Klinikum zeitlich nach Covid-19-Impfungen gestorben seien. Für die Behauptungen gibt es keinerlei Belege. Das Klinikum weist sie als Falschmeldung zurück, das Landgericht Hamburg hat eine einstweilige Verfügung gegen die Verbreiterin des Gerüchts erlassen.

(a) Fact check on Corona vaccines, Wave I

Thema Impfstoffhersteller

Fakten-Check

dpa • factchecking

Ungenauere Behauptungen über Impfstoffhersteller

23/05/2021, 02:51 PM (CEST)

Die Impfstoffe gegen das Coronavirus Sars-CoV-2 werden seit rund einem halben Jahr in vielen Ländern der Welt eingesetzt. Eine Übersicht auf Facebook stellt nun einige Behauptungen über die Hersteller der bislang in der EU zugelassenen Impfstoffe auf ([hier archiviert](#)). Was ist dran an den Warnungen?

Bewertung

Einige der Aussagen sind ungenau oder aber es fehlt wichtiger Kontext. Zum Beispiel wird in den meisten europäischen Ländern mit dem Corona-Impfstoff von AstraZeneca geimpft.

Fakten

Über den Impfstoff des Unternehmens **AstraZeneca** heißt es, dass er von mehr als 20 europäischen Ländern wegen schwerer und tödlicher Nebenwirkungen «suspendiert» worden sei. Richtig ist, dass viele Länder den Einsatz des Impfstoffs im März vorübergehend ausgesetzt hatten oder haben. Allerdings verzichteten nur **Dänemark und Norwegen** dauerhaft auf den Impfstoff. Die meisten anderen europäischen Staaten setzen ihn inzwischen wieder ein, wenn auch meist eingeschränkt. In Deutschland und Italien etwa wird der Impfstoff von den zuständigen Stellen für Menschen über 60 empfohlen.

Über die Firma **Johnson & Johnson** wird in dem Instagram-Beitrag behauptet, dass sie «in Hunderttausenden Prozessen verurteilt» worden sei. Für eine derart hohe Zahl gibt es keine Belege. Richtig ist, dass Johnson & Johnson unter anderem wegen eines Skandals um Asbest in Baby puder mit vielen Klagen konfrontiert ist.

Dem Unternehmen **Moderna** wird unterstellt, dass keiner seiner Impfstoffe in Zulassungsstudien die sogenannte Phase 3 überstanden habe. Durch die Formulierung in der Vergangenheitsform wird ein Eindruck erweckt, als ob die Studien abgebrochen worden seien oder die Behörden eine Zulassung verweigert hätten. Dem ist jedoch nicht so: Ein Blick in eine offizielle Übersicht der US-Behörden zeigt, dass Studien zu anderen Moderna-Impfstoffen derzeit zum Teil noch laufen und sich noch in anderen Phasen befinden. Die Studien laufen hier lediglich deutlich länger als beim Corona-Impfstoff des Unternehmens.

Dessen Zulassung wurde zwar beschleunigt, in Europa etwa durch das sogenannte Rolling-Review-Verfahren. Aber auch dieser Impfstoff wurde vor der Zulassung mehr als 15 000 Testpersonen gespritzt, eine Größenordnung, die der vieler anderer Phase-3-Studien entspricht. Dabei wurde eine sehr hohe Wirksamkeit gegen Sars-CoV-2 festgestellt, die die Risiken deutlich übersteigt.

(c) Fact check on Corona vaccines, Wave II

Thema Radioaktivität

Fakten-Check

AFP • Faktencheck

Diese Geschichte über radioaktive Corona-Impfstoffe ist frei erfunden

Mehrere Hundert Facebook-User haben seit Mitte Dezember die Behauptung einer Person geteilt, die sich als Angestellte des Bundesinstituts für Arzneimittel und Medizinprodukte (BfArM) ausgibt. Diese will eine Ampulle eines Impfstoffes aus dem Institut gestohlen und diese zu Hause untersucht haben. Ergebnis: Der Impfstoff soll angeblich leicht radioaktiv sein. Das BfArM lagert aber keinen Impfstoff, den Angestellte stehlen könnten. Das Institut ist nicht für Impfstoff-Zulassungen zuständig.

(b) Fact check on Corona vaccines, Wave I

Thema Impfpampulle

Fakten-Check

dpa • factchecking

Zahl auf Impfpampulle fehlinterpretiert: kein Herstellungsdatum

03/03/2021, 06:26 PM (CET)

Eine Zahlenkombination auf einer Impfdosis soll angeblich beweisen, dass der Impfstoff schon lange vor Beginn der Corona-Pandemie entwickelt worden sein soll. Konkret geht es um ein Foto eines Fläschchens, das den Impfstoff der Firma AstraZeneca enthält ([hier archiviert](#)). Dieses Foto verbreitet sich auf Facebook. Auf dem Etikett ist der Name «AstraZeneca» zu sehen, außerdem «COVID-19 Vacc» und drei Zahlenkombinationen: «AB0003», «05-2021» und senkrecht an der Seite des Etiketts die Nummer «3041572008». Letztere wird in dem zum Foto gehörigen Beitrag als Chargennummer interpretiert. Angeblich soll sie verraten, wo und wann das Produkt hergestellt worden sei, wer es produziert habe sowie welche Maschinen und Inhaltsstoffe dabei zum Einsatz gekommen seien. Im Beitrag zu dem Foto heißt es zur Interpretation des Etiketts: «Herstellungsdatum : 2008» und «Verfallsdatum : Mai / 2021».

Die seitlich stehende Nummer «3041572008» ist also nicht die Chargennummer. Auf dpa-Anfrage teilte AstraZeneca mit, dass es sich dabei um einen Komponenten-Code handelt, der für die jeweilige Produktionsstätte feststehend sei und für die Rückverfolgung in internen Systemen verwendet werde. Auf einen früheren Produktionsbeginn des Impfstoffs deute die Zahl nicht hin. «Die kommerzielle Produktion des Covid-19-Impfstoffs von AstraZeneca fing nicht vor Herbst 2020 an», teilte eine Sprecherin mit.

(d) Fact check on Corona vaccines, Wave II

Figure A.5: Fact checks on Corona vaccines.

Thema Proteinbedarf

Fakten-Check



Deutsche Gesellschaft für Ernährung e.V.
Der Wissenschaft verpflichtet – Ihr Partner für Essen und Trinken

14. Gibt es gesonderte Empfehlungen von der DGE für die Proteinzufuhr für Sportler?

Für erwachsene Breitensportler*innen (4–5 Mal je Woche 30 Minuten körperliche Aktivität bei mittlerer Intensität) gibt es keine gesonderte Empfehlung. Zur Sicherstellung der Proteinversorgung reicht eine Zufuhr in Höhe der empfohlenen Zufuhr von 0,8 g Protein/kg Körpergewicht pro Tag aus.

Im ambitionierten Breitensport und Leistungssport (mind. 5 Stunden Training pro Woche) kann eine sportart- und belastungsspezifisch angepasste Proteinzufuhr den Trainingsprozess sinnvoll unterstützen und die Leistungsbereitschaft fördern. Über Zufuhrmenge, Art der Proteinquelle, optimale Aminosäurezusammensetzung sowie Zeitpunkt der Zufuhr wird teilweise kontrovers diskutiert. Die *International Society of Sports Nutrition* und das *American College of Sports Medicine* empfehlen je nach Sportart und Trainingsziel, -intensität, -umfang oder Wettkampfphase eine flexibel angepasste Proteinversorgung mit ca. 1,2–2,0 g/kg Körpergewicht pro Tag. Die Proteine sollten über den Tag verteilt und im Rahmen von Mahlzeiten und nicht als Supplemente zugeführt werden.

Im Positionspapier der Arbeitsgruppe Sporternährung der DGE zu Proteinzufuhr im Sport werden aktuelle Erkenntnisse zu physiologischen Wirkungen der Proteinzufuhr im Sport, unter besonderer Berücksichtigung des Zufuhrmenge- bzw. des Dosis-Wirkung-Aspekts, dargestellt.

(a) Fact check on Nutrition, Wave I

Thema Vitamin B12

Fakten-Check

verbraucherzentrale

Was steckt hinter der Werbung zu Vitamin B12?

Immer wieder hört man von einem weit verbreiteten Vitamin B12-Mangel in der Bevölkerung. Wissenschaftliche Belege für diese Behauptungen gibt es keine. Sowohl Berechnungen zur Vitamin B12-Versorgung als auch Untersuchungen im Blut zeigten, dass ein Vitamin B12-Mangel selten ist. In der Regel führt die westliche Ernährungsweise mit einem hohen Anteil an tierischen Lebensmitteln eher zu einer Vitamin B12-Übersorgung.

(c) Fact check on Nutrition, Wave II

Thema Vitamin C

Fakten-Check



Deutsche Gesellschaft für Ernährung e.V.
Der Wissenschaft verpflichtet – Ihr Partner für Essen und Trinken

3. Gibt es hierzulande einen Vitamin-C-Mangel?

In industrialisierten Ländern kommen Vitamin-C-Mangelzustände praktisch nicht mehr vor. Klassische klinische Vitamin-C-Mangelzustände sind beim Säugling die Moeller-Barlow-Krankheit und beim Erwachsenen der Skorbut (früher oft als „Seefahrerkrankheit“ beschrieben). Dabei sind die Knochenbildung und das Wachstum beim Säugling und Kind gestört. In späteren Lebensabschnitten sind die Symptome schlechte Wundheilung, Gelenkschmerzen, Infektionen, Neigung zu Blutungen in der Haut, den Schleimhäuten, der Muskulatur und den inneren Organen sowie Zahnausfall. Diese Störungen treten bei Erwachsenen nur bei dauerhaft fehlender Vitamin-C-Zufuhr auf. Bereits 10 mg Vitamin C pro Tag verhindern Skorbut.

(b) Fact check on Nutrition, Wave I

Thema Vitamin E

Fakten-Check



Wie viel Vitamin E brauchen wir?

Der genaue Bedarfswert ist nicht bekannt. Der Schätzwert für eine angemessene Zufuhr für Erwachsene (25 bis < 51 Jahre) liegt pro Tag nach den D-A-CH-Referenzwerten bei 12 mg (Frauen) und 14 mg Vitamin E (Tocopherol-Äquivalente). Für **schwangere Frauen** liegt der Schätzwert für eine angemessene Zufuhr bei 13 mg und für **Stillende** bei 17 mg pro Tag.

Näheres zu allen Altersklassen bzw. Personengruppen sowie Geschlecht erfahren Sie in den [D-A-CH-Referenzwerten](#). Weitere Informationen erhalten Sie unter [Deckung des Tagesbedarfs an Vitaminen](#).

Zu viel/zu wenig Vitamin E?

- **Ein Zuviel** an Vitamin E ist selten, dennoch sollten hohe Dosierungen (Supplemente) über einen längeren Zeitraum vermieden werden, da sie u.a. Magen-Darm-Probleme und ein erhöhtes Blutungsrisiko verursachen können. Als tolerierbare Gesamtzufuhr gelten 300 mg pro Tag (Tolerable Upper Intake Level / EFSA).

(d) Fact check on Nutrition, Wave II

Figure A.6: Fact checks on Nutrition.

A.4. (Un-)trustworthy areas in fakes and facts

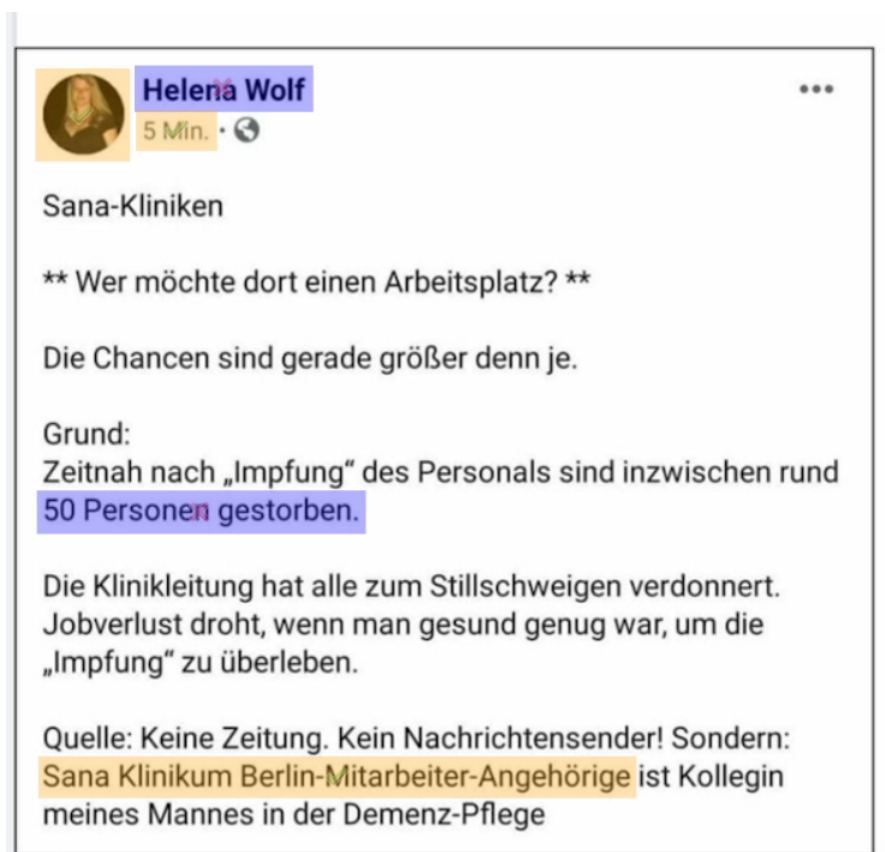


Figure A.7: Exemplary fake from the ex-post survey, where three elements are marked as “trust-worthy” (in orange) and two elements are marked as “untrustworthy” (in blue).

A.5. Translations of the images

This section provides translations for the images in Sections A.3 and A.4.

A.5.1. Figure A.1

Figure A.1 (a) Topic on sana-clinics

Heike Tamm:

“So much on that. The nursing staff is becoming scarce, it’s dying.”

Helena Wolf:

Sana-clinics

** Who wants a job there? **

The chances are better than ever.

Reason:

Briefly after “vaccination” of staff members, 50 people have died by now.

The clinics management forced everyone to remain silent.

Risking losing one's job, if one was healthy enough to survive the “vaccination”.

Source: No newspaper. Now news station! But: A Sana clinic Berlin-staff member-family member is a colleague of my husbands in dementia-care.

Figure A.1 (b) Topic on radioactivity

Christian I forgot:

I work at BfArM (federal institute for drugs and medical products) and I had the opportunity to smuggle an ampoule of the vaccine secretly out of the building to analyse it at home. I discovered that this alleged vaccine glows in the dark and is slightly radioactive.

At home I measured approximately 70 microsievert per hour. This is enough to cause diseases as well as genetic damages.

There might be radium in this vaccine.

Alert everyone about this madness !!!

Here is a picture of the "vaccine" that I pulled from the ampoule with an injection needle:
(picture)

Figure A.1 (c) Topic on vaccine ampoules

Mary Stingelin:

Interesting!

Marion Rafaela Tillmann:

OPEN YOUR EYES WHEN BUYING VIRUSES (eye emoji, winking emoji)

Otherwise even the dumbest would see the fake!! The vaccine was already produced in 2008 and expires this year in May???

Where does a batch number duty exist?

An obligation to a batch number exists, whenever dealing with medical devices or drugs.

In case of an inquiry, the batch number provides:

Where the product was produced

When the product was produced

Who produced it

Which machines were used for production

And finally, which raw materials and ingredients it contains

If this is true??

Production date: 2008

Expiration date: May/2021

(in the picture:

Vakzin from AstraZeneca

Vaccine-delivery to states expe[cted]

10.485 new infections and 689 new deaths. Corona-pandemic: Demands [...])

Figure A.1 (d) Topic on producers of vaccine

Fred Walter:

(in the picture:

CHOOSE YOUR COVID-! VACCINE)

Facts you all should know about...

Pfizer: 4.7 billion penalty for false allegations, violation of the safety of drugs and medical devices, off-label-promotion, corrupt practices, bribery and bribe money.

Moderna: Hasn't brought a vaccine onto the market since its foundation, even though there were more than nine candidates. None of those survived the clinical phase 3.

Johnson & Johnson: Convicted in hundred thousand processes because of toxic and/or dangerous products. Including drugs, shampoos, medical devices as well as baby powder polluted with asbestos!

AstraZeneca: Was suspended by 20+ European countries because of severe and deadly side effects!

No worries, you are in save hands!

Once you are vaccinated, remember social distancing and to wear a mask.

As you can still spread COVID-19. And trust the experts!

A.5.2. Figure A.2

Figure A.2 (a) Topic on protein demand

Clever fit Ludwigsfelde is here: clever fit Ludwigsfelde

Do you achieve your daily demand of protein?

Protein is by far the most important macronutrient, because how are new muscles supposed to grow, if there is no construction material?

The general advice is: a daily dose of 2g per kg body weight

Do you achieve to meet your daily protein demand?

(Hashtags)

Figure A.2 (b) Topic on Vitamin C

Îyadh Chaouachi:

(picture already in English)

Figure A.2 (c) Topic on Vitamin B12

Dr. Rath health-foundation:

Up to 40 percent of adults show a shortage of vitamin B12, which has various consequences.

ISSUU.COM

DR. RATH HEALTH LETTER – version 09/21 – June 2021 – VITAMIN B12: Shortage can cause tinnitus

Up to 40 Percent of adults show a shortage of vitamin B12, which has various consequences. If there is a lack of Vitamin B12, the vessel and the nervous system of the ear can be damaged for example and a tinnitus might be caused. The symptoms are diffuse and the course of disease often insidious – d. . .

Figure A.2 (d) Topic on Vitamin E

Îyadh Shawashi:

Those, who often eat outside, should probably supplement at least 30 mg natural vitamin E (D-Alpha) (resp. 90mg synthetic, equalling 60 aTE/200ie as DL-alpha), as the high concentration of Omega-6 in plant based oils consumes extreme amounts of vitamin E.

See how in the following graphic the introduction of 30g and then 60g seed oil presses down the vitamin E concentration notably.

Without supplementation it is just a catastrophe with the high amounts of Omega-6 in modern diets

(From mark “NONE”)

[Graphic from Herwitt 1960]

A.5.3. Figure A.3

Figure A.3 (a) Topic on Sputnik V

ZDF today:

Even before a possible permit in the EU, Bavaria secured more than two million doses of the Russian vaccine Sputnik V for itself: (Link)

(In the picture: Bavaria secures 2.5 million doses of Sputnik V for itself)

Figure A.3 (b) Topic on Effectiveness

ZDF today:

The vaccine from Biontech/Pfizer protects according to manufacturer information adolescence responsibly against a Covid-19-infection, too. Based on the phase-3-study the effectiveness in adolescence is labelled with a hundred percent.

(in the picture: Breaking news

New study: Biontech-vaccine shows high effectiveness in adolescence)

ZDF.DE

Biontech announces a hundred percent effectiveness in adolescence.

Figure A.3 (c) Topic on Staying Away

FAZ.NET – Frankfurt general newspaper:

In Saarland more than 100 missed vaccination appointments of medical staff cause furore:

(Picture)

FAZ.NET— FROM FAZ.NET – Frankfurt general newspaper Why do more than 100 doctors not get vaccinated?

Figure A.3 (d) Topic on vaccination complications

Rheinische Post:

Reasons seem to be complications with vaccination.

(Picture)

RP-ONLINE.DE

18 vaccinees want to take the state NRW into account

Because of complications, 18 vaccinees currently want to take the state NRW into accou[nt]

A.5.4. Figure A.4

Figure A.4 (a) Topic on plastic

AOK Bavaria:

Life in plastic ... isn't fantastic! (littering emoji)

Did you test once how much of an example you are, too, and inspected your apartment on accruing plastic?

You will probably be as shocked as Inka. (shocked emoji)

Editor's note: On average we consume approximately 5g micro plastic per week.

You can find tips on zero waste shopping here: (link)

(In the picture: Plastic is everywhere

Click to zoom

#ItStartsWithYou)

Figure A.4 (b) Topic on sugar

AOK Bavaria:

So sweet, that one could turn sour: Children are eating way too much sugar — but they are not alone. Current studies show, that adults, too, have fallen excessively for the sweet stuff.

Due to that 40% of women and 30% of men are eating too much sugar.

We need to talk about that. The right frame for it: The 3rd German sugar-reduction-summit on 27.10.2020. Here, guests like Julia Klöckner or Renate Künast will report on the national reduction strategy. Their ... Load more

(In the picture: Children consume 75% too much sugar)

Figure A.4 (c) Topic on moving

Federal institute for nutrition:

80 percent of adolescents are moving less than recommended by the world health organisation!

The network “Healthy into life — network young family” as well as IN SHAPE — Germanys initiative for healthy nutrition and more movement provide various information, recommendations, material and recipes on their websites

(picture)

BZfE.DE

Less Couch-Potatos –BZfE

Federal institute for nutrition (BZfE): The competence- and commu. . .

Figure A.4 (d) Topic on nutrition scores

Upfit:

Time for a food comparison (boxing glove emoji):

Do you regularly eat chia or linseeds? And if so, with what do you like them best? (smiley emoji)

Chia seeds are the Latin American pendant to the European linseeds. The nutrition scores of both of them don’t vary much. Linseeds have 471 kcal for every 100g, chia seeds 475 kcal.

Chia and linseeds deliver high-value protein and omega-3- and omega-6-fatty acids. (strong arm emoji)

The especially high share of calcium and vitamin A in the chia seeds sets those apart. Linseeds meanwhile contain a lot of potassium and iron and zinc are contained as well.

All in all, both seeds provide important nutritions and are very healthy. Those who prefer a regional origin of their food should choose the linseeds. These are also cheaper than the chia seeds most of the time.

#projectlivehealthy #fatlogicfree #eatingbetter #looseweightgermany #dietarydiary #easy-cooking #eatinghealthy #thoughtfuleating #looseweightwithfun #youarewhatyoueat #looseweight-intuitively #lowincaloriesrecipes

(In the picture: food comparison

Chia seeds — line seeds

Learn more on:

UPFIT)

A.5.5. Figure A.5

Figure A.5 (a) Topic on Sana-clinics

Fact-check

CORRECTIV

Research for the society

Rumor about mortality after Corona-vaccination in Sana-clinic is fictitious.

Especially on Facebook and Telegram there currently is a rumor about the death of 50 staff members of Berlin's Sana-clinic shortly after Covid-19-vaccinations. There is no evidence for these claims. The clinic dismisses them as false reports, the state court of Hamburg has issued an injunction against the propagator of this rumor.

Figure A.5 (b) Topic on radioactivity

Fact-check

AFP fact check

This story about radioactive Corona-vaccinations is fictitious.

Several hundred Facebook-User shared the claim of a person, who impersonated a staff member of the federal institute for drugs and medical devices (BfArM), since mid-December. This person claims to have stolen an ampoule of the vaccine from the institute and to have tested it at home.

Result: The vaccine is supposedly slightly radioactive. The BfArM however isn't storing any vaccines staff members could steal. The institute isn't responsible for vaccine-permits.

Figure A.5 (c) Topic on vaccine manufacturer

Fact-check

dpa factchecking

Uncertain claim about vaccine manufacturer

The vaccines against the Coronavirus Sars-CoV-2 are being used for about six months in many countries around the world. An overview on Facebook now raises some claims about the manufacturers of the vaccines that are authorized in the EU up to now (archived here). What's up with the warnings?

Evaluation

Some of the statements are vague or there is important context missing. For example most European countries use the vaccine from Astrazeneca to vaccinate.

Facts

About the vaccine of the company Astrazeneca is said that it is suspended by more than 20 European countries because of severe or deadly side effects. True is, that many countries paused or have paused the use of the vaccine temporarily in March. However, only Denmark and Norway renounced the vaccine permanently. Most of the other European states are already using the vaccine again, even though mostly with some restrictions. In Germany and Italy the vaccine is additionally only recommended to people over 60 by the responsible bodies.

About the company Johnson & Johnson the Instagram-post states, that it "was convicted in hundred thousand processes". There is no proof for a number this high. It is true, that Johnson & Johnson is among other things confronted with many complaints because of a scandal about asbestos in baby powder.

It is assumed that none of the vaccines from the company Moderna surpassed the so-called phase 3 of their permission-studies. Due to the formulation in past tense the impression arises

that the study was cancelled or that the authorities denied the permit. It is not like that though: A glance into an official overview by the US-authorities shows, that studies on other Moderna-vaccinations are currently still running or are still in other phases. The studies are just running a lot longer here than the companies Corona-vaccine.

This permit was indeed accelerated, in Europe for example by the rolling-review-procedure. Nevertheless more than 15 000 test subjects, a size similar to many other phase-3-studies, were injected with this vaccine before permission. With this a large effectiveness against Sars-CoV-2 was determined, which exceeds the risks by far.

Figure A.5 (d) Topic on vaccine ampoule

Fact-check

dpa factchecking

Number on vaccination ampoule misinterpreted: no production date

A combination of numbers on a vaccination dose supposedly proves that the vaccine has been developed long before the Corona-pandemic. In detail this is about a picture of a small flask that contains the vaccine of the company Astrazeneca (archived here). This picture spread on Facebook. On the label the name “AstraZeneca” is depicted, additionally “COVID-19 Vacci” and three combination of numbers: “AB0003”, “05-2021” and vertically on the side of the label the number “3041572008”. The last is interpreted as batch number in the post concerning the picture.

It is supposed to reveal, where and when the product was produced, who produced it and as which machines and ingredients have been used. In the post with the picture it was said: “production date: 2008” and “expiration date: May /2021”

The vertical number “3041572008” on the side is not the batch number. In response to dpa’s request, Astrazeneca replied that the number series is a component-code, which is set for each production site and is needed in internal systems for trace back. The number is not referring to an earlier start of production. “the start of commercial production of the Covid-19-vaccine from Astrazeneca was not before fall 2020”, a speaker announced.

A.5.6. Figure A.6

Figure A.6 (a) Topic on protein demand

Fact-check

German society of nutrition e.V. (e. V. = enlisted association)

Committed to science – your partner for food and drinks

14. Are the special recommendations about protein intake for athletes by the German society of nutrition (DGE)?

For grown up general athletes (30 minutes of physical activity with medium intensity 4-5 times a week) there are no special recommendations. To ensure meeting ones protein demands, the recommended amount of 0.8 g protein per kg body weight per day is enough.

In ambitious mass sport and competitive sports (minimum of 5 hours training per week) a protein intake that is adjusted specifically according to the kind and load of sport, might support the process of training and promote the willingness to perform. There is a partly controversial debate about the amount of intake, type of protein source, optimal aminoacid composure as well as time of intake. The International Society of Sports nutrition and the American College of Sports Medicine recommend flexibly adjusted protein supply based on type of sport and goal, intensity and load of training with approximately 1.2 – 2.0 g/kg body weight per day. The proteins should be evenly distributed over the day and be consummated during meals and not as supplements.

In the policy paper of the working group on sports nutrition of the DGE about protein intake in sports current findings about psychological effects of protein intake in sports, with special consideration of the amount of intake- and dose-effect-aspects, are being discussed.

Figure A.6 (b) Topic on Vitamin C

Fact-check

German society of nutrition e.V. (e. V. = enlisted association)

Committed to science – your partner for food and drinks

3. Is there a vitamin-C-lack in this country?

In industrialised nations vitamin-C-lacks barely exist anymore.

Common clinical vitamin-C-lacks are the Moeller-Barlow disease in infants and the Skorbut (also known as the “sailor’s disease”) in adults. With these diseases the formation of new bones and growth in general is disturbed in infants and children. Later on the symptoms are bad healing of wounds, joint pain, infections, tendency to bleeding within the skin, the mucosae, the muscles and inner organs, as well as losing one’s teeth. In adults the disorders only appear if there is a continuous lack in vitamin-C supply. 10 mg vitamin C a day can already stop Skorbut.

Figure A.6 (c) Topic on Vitamin B12

Fact-check

User headquarters

What’s behind the advertisement for vitamin B12?

Regularly one hears about a widespread vitamin B12-lack in general population. There is no scientific proof for this claim. Calculations about the vitamin B12 supply as well as examinations of blood showed, that a lack of vitamin B12 is rare. Usually the western diet with a large portion of animal products rather causes a vitamin B12 overflow.

Figure A.6 (d) Topic on Vitamin E

Fact-check

health.GV.AT

open health portal Austria

How much vitamin E do we need?

The exact amount is unknown. An estimate for adults (25 till ; 51 years) is according to D-A-CH-reference values 12 mg of vitamin E (Tocopherol-equivalents) a day for women and 14 mg for men. For pregnant women an appropriate estimate is 13 mg and 17 mg a day for nursing women.

You can find more about all age-levels and groups of people as well as genders in the D-A-CH-reference values. More information under achieving the daily demand of vitamins.

Too much/not enough vitamin E?

A too much vitamin E is rare, nevertheless high doses (supplements) over a longer period of time should be avoided, as they might cause gastrointestinal problems and a heightened risk of bleeding. About 300mg per day are considered a tolerable upper intake level.

A.5.7. Figure A.7

Figure A.7

Helena Wolf

Sana-clinics

** Who wants a job there? **

The chances are better than ever.

Reason: Briefly after “vaccination” of staff members, 50 people have died by now.

The clinics management forced everyone to remain silent. Risking loosing ones job, if one was healthy enough to survive the “vaccination”.

Source: No newspaper. Now news station! But: A Sana clinic Berlin-staff member-family member is a colleague of my husbands in dementia-care.

B. Hypotheses

We pre-registered the following hypotheses in the AEA Registry under registry number AEARCTR-0008199:

Hypothesis 1a: *In the short-run, the fact checking intervention reduces the credibility of and increases factual knowledge about the corrected “fake news” as compared to participants without intervention.*

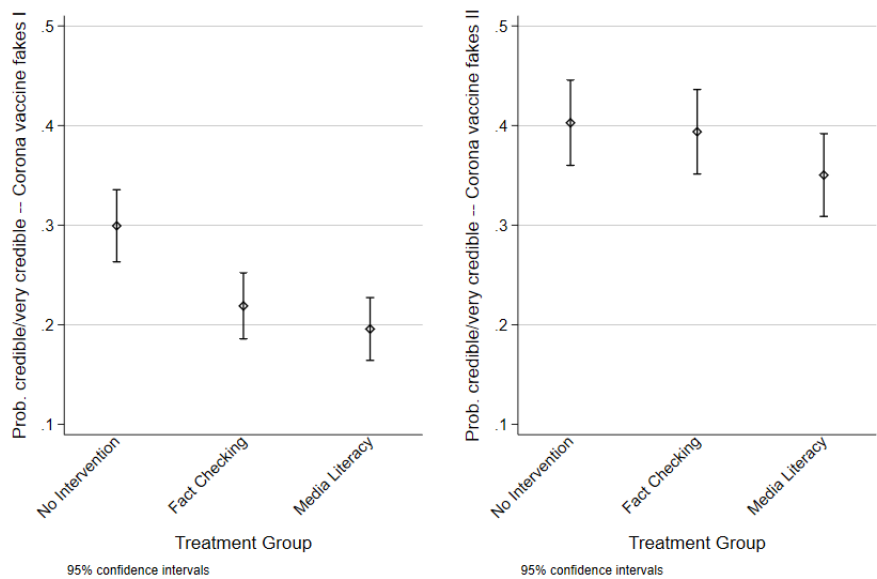
Hypothesis 1b: *In the short-run, participants who received the fact checking intervention are more likely to state that they are willing to get vaccinated against Covid-19 than participants without intervention.*

Hypothesis 2a: *In the short- and in the longer-run, the media literacy intervention reduces the credibility of and increases factual knowledge about all “fake news” as compared to participants without intervention.*

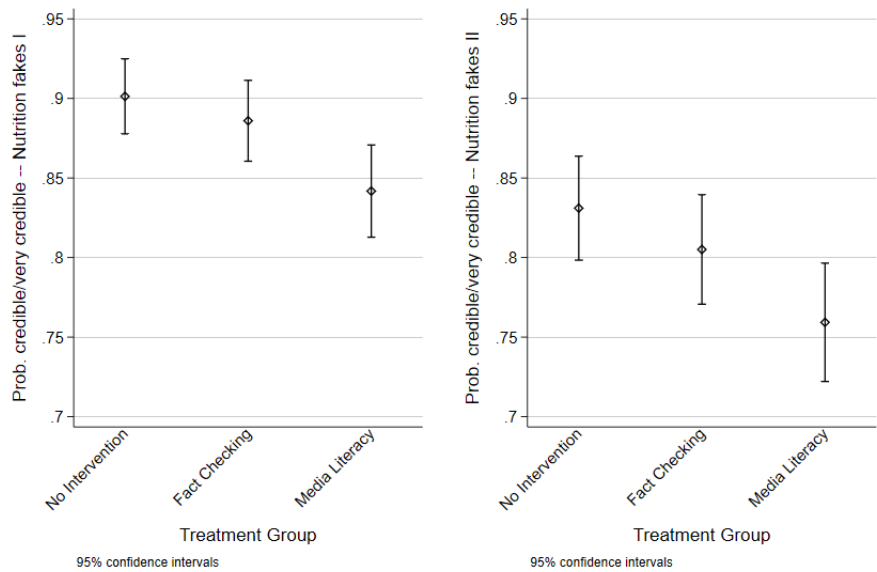
Hypothesis 2b: *In the short- and in the longer-run, participants who received the media literacy intervention are more likely to state that they are willing to get vaccinated against Covid-19 and abstain from unnecessary dietary supplements than participants without intervention.*

Hypothesis 2c: *The longer-term effects of the media literacy intervention are smaller than its short-term effects.*

C. Omitted figures

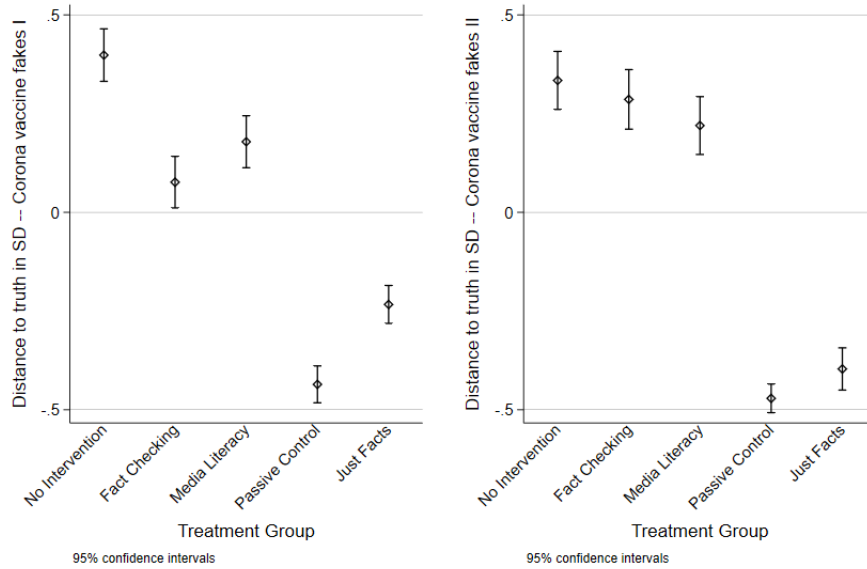


(a) Average credibility of **Corona fakes** per treatment group and survey wave

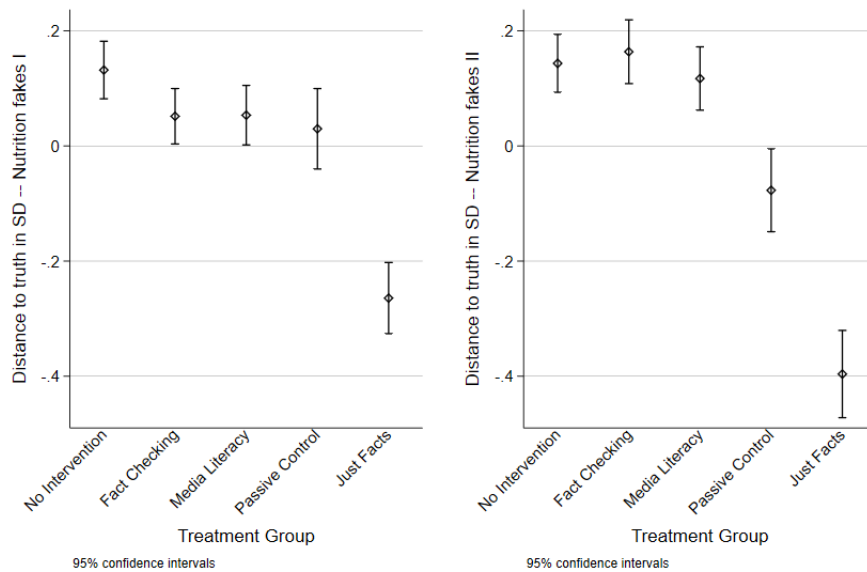


(b) Average credibility of **nutrition fakes** per treatment group and survey wave

Figure A.8: Average credibility of fakes.

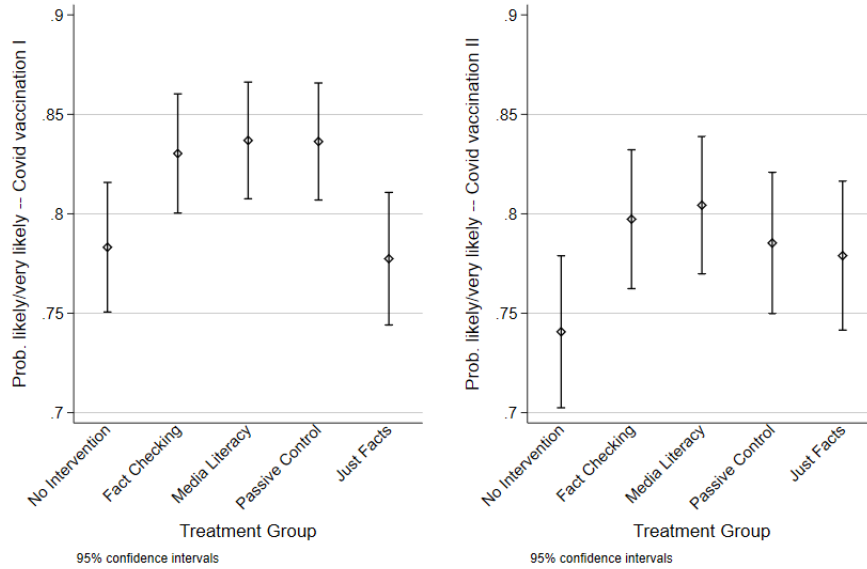


(a) Average standardized distance to the correct answers to questions on **Corona fakes** per treatment group and survey wave.

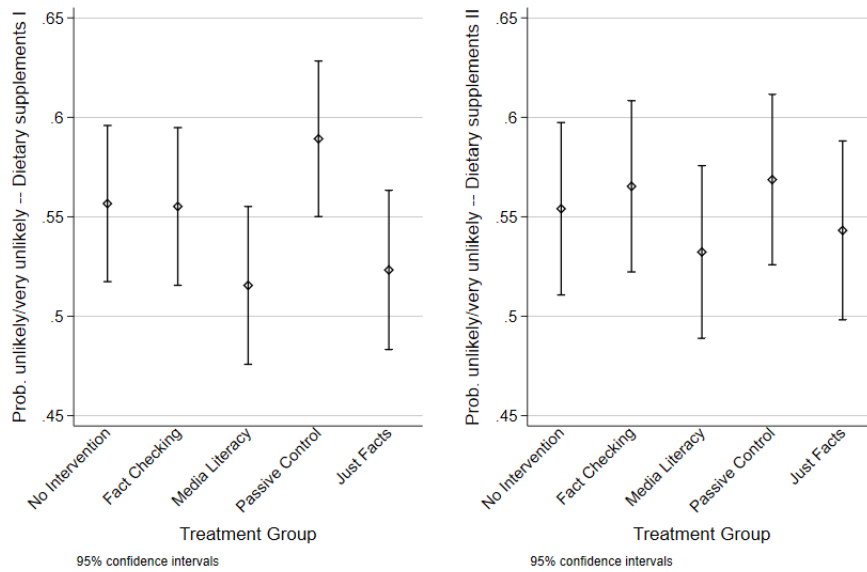


(b) Average standardized distance to the correct answers to questions on **nutrition fakes** per treatment group and survey wave.

Figure A.9: Average standardized distance to the correct answers to the factual knowledge questions on fakes.

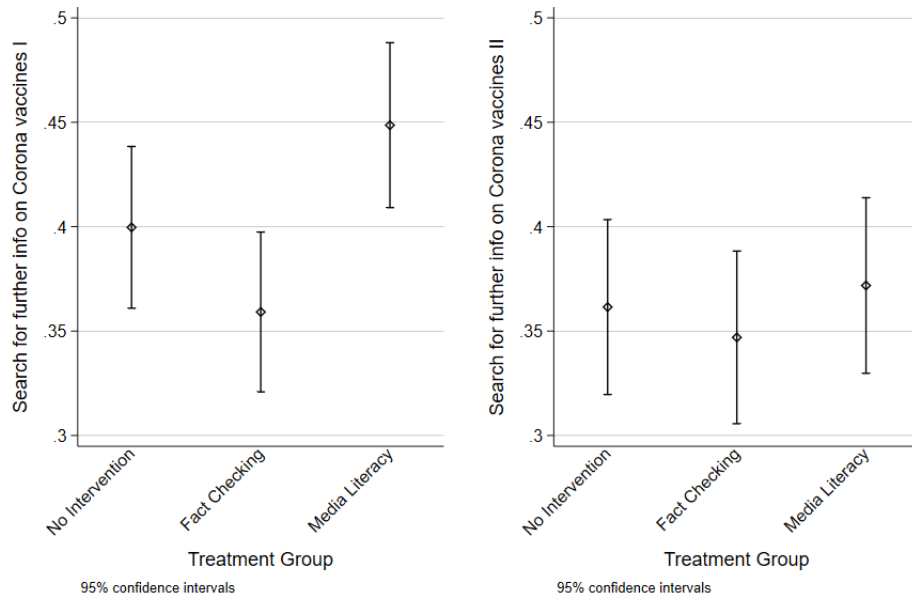


(a) Average reported probability to get vaccinated or boosted against Corona per treatment group and survey wave.

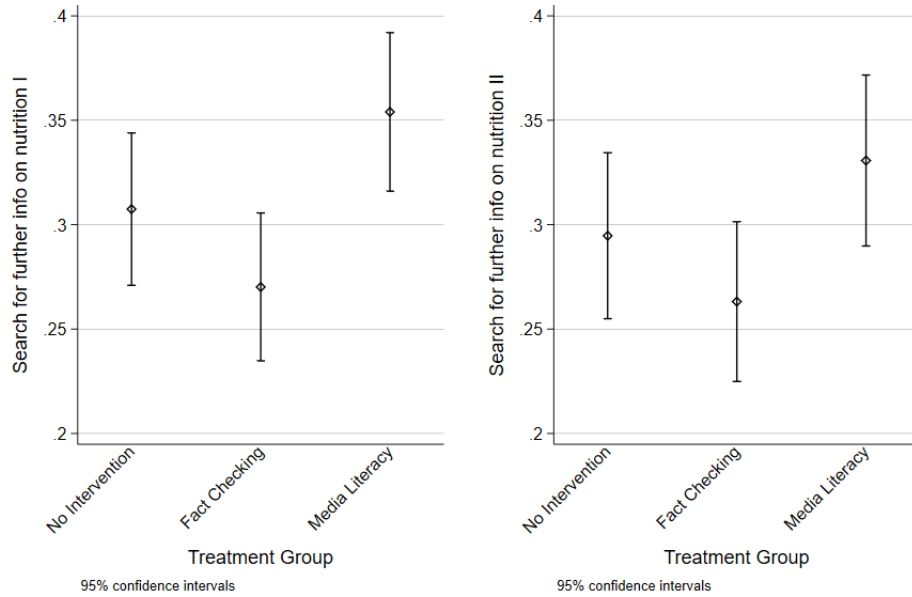


(b) Average reported probability to consume dietary supplements per treatment group and survey wave.

Figure A.10: Average reported probability to get vaccinated or boosted against Covid-19 and to consume dietary supplements.

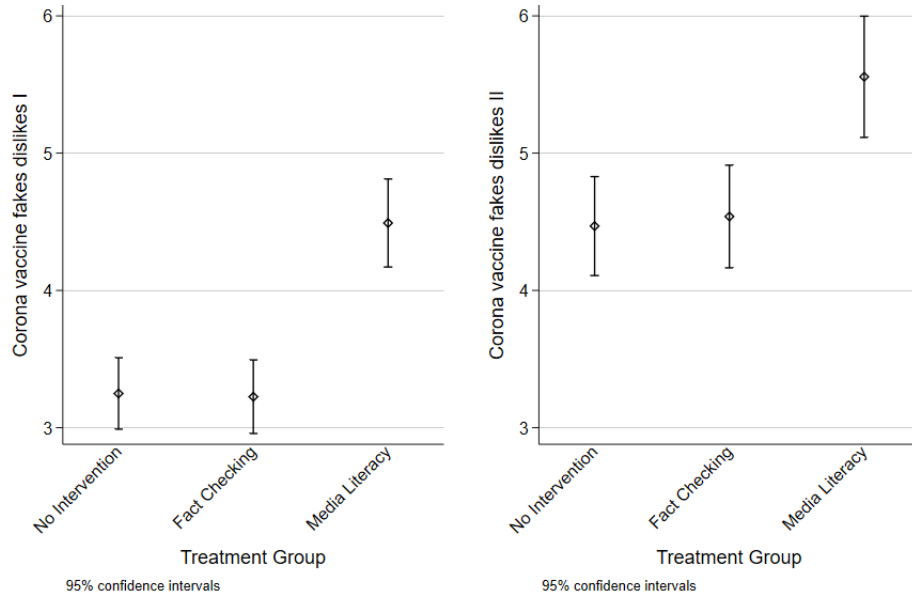


(a) Share of participants who searched for further information on Corona vaccines.

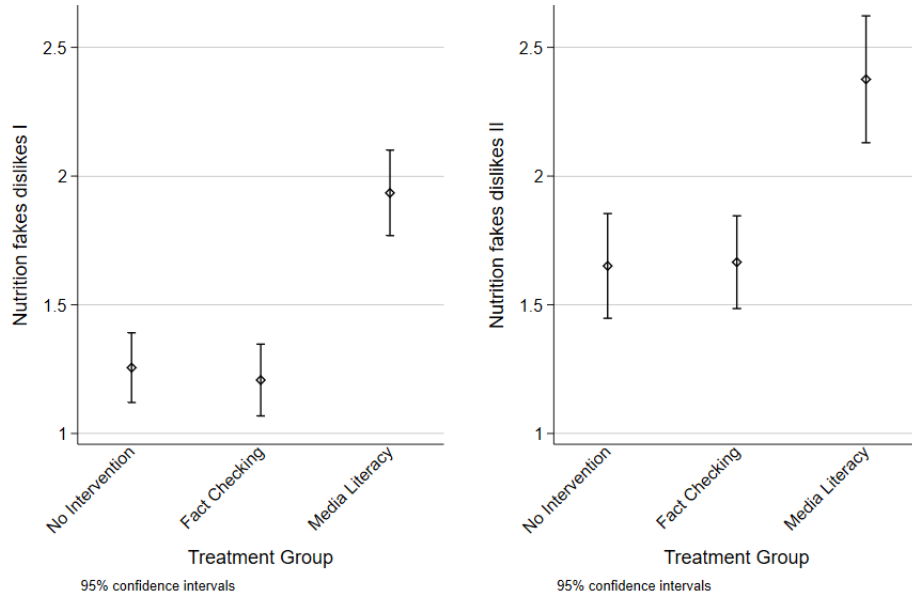


(b) Share of participants who searched for further information on nutrition.

Figure A.11: Share of participants who searched for further information online.

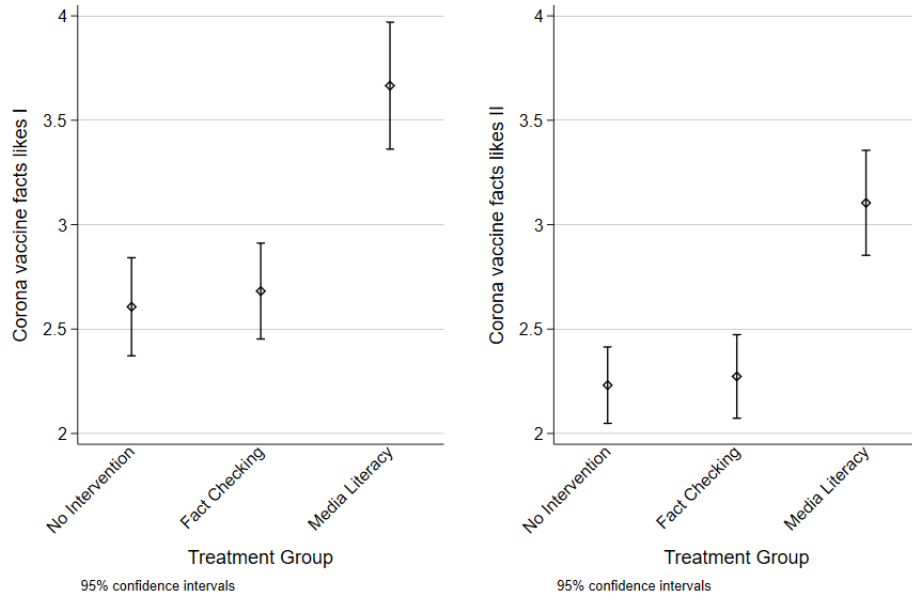


(a) Average number of dislikes for fakes on Corona vaccines.

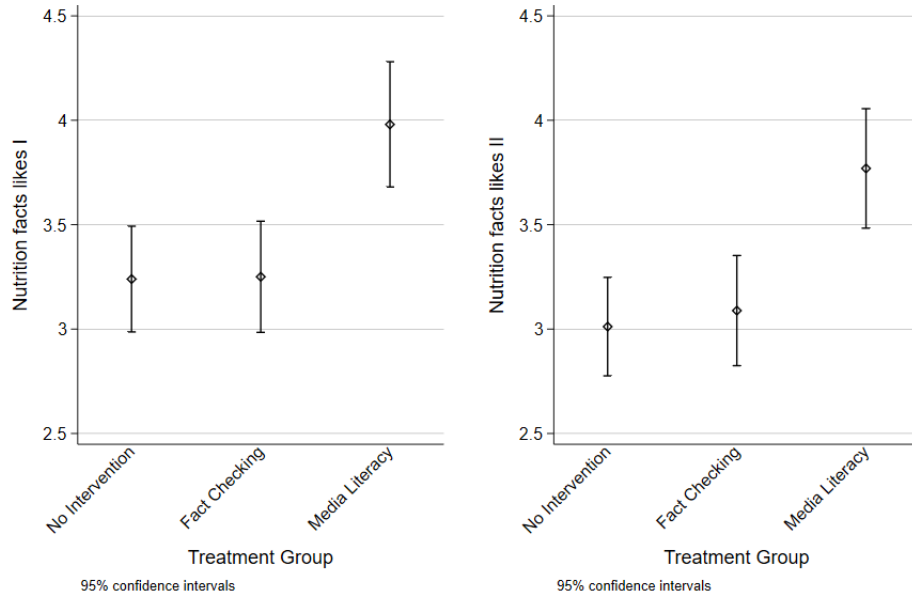


(b) Average number of dislikes for fakes on nutrition.

Figure A.12: Average number of dislikes for fakes.

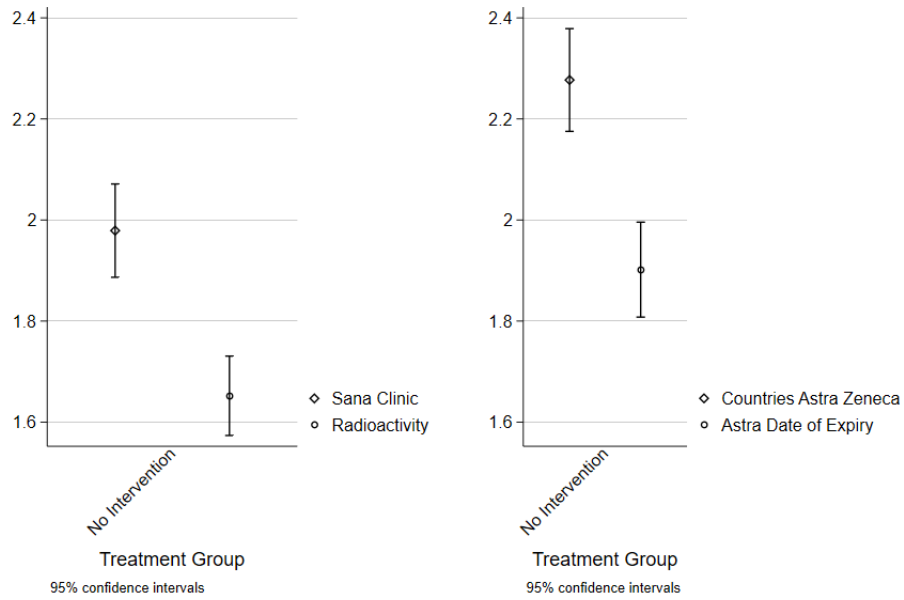


(a) Average number of likes for facts on Corona vaccines.

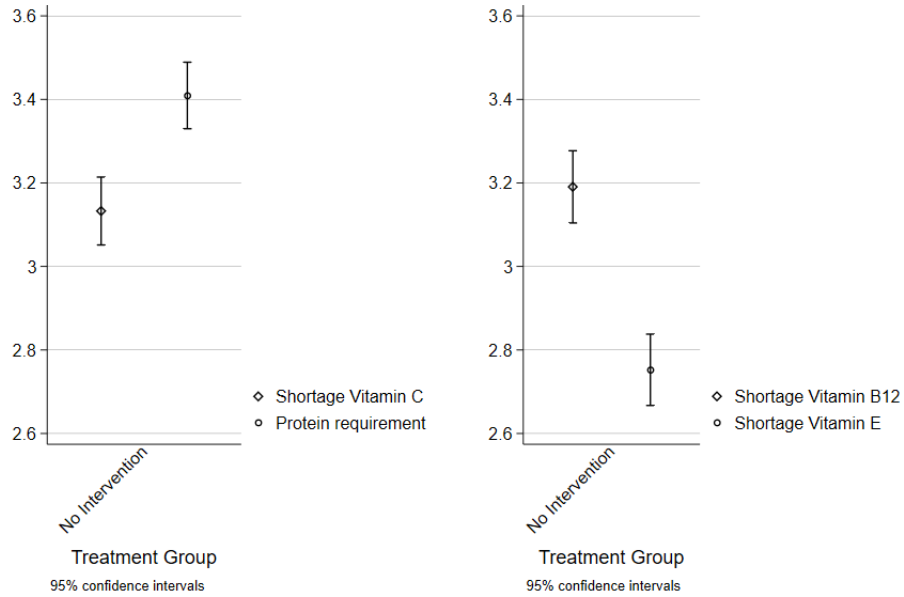


(b) Average number of likes for facts on nutrition.

Figure A.13: Average number of likes for facts.

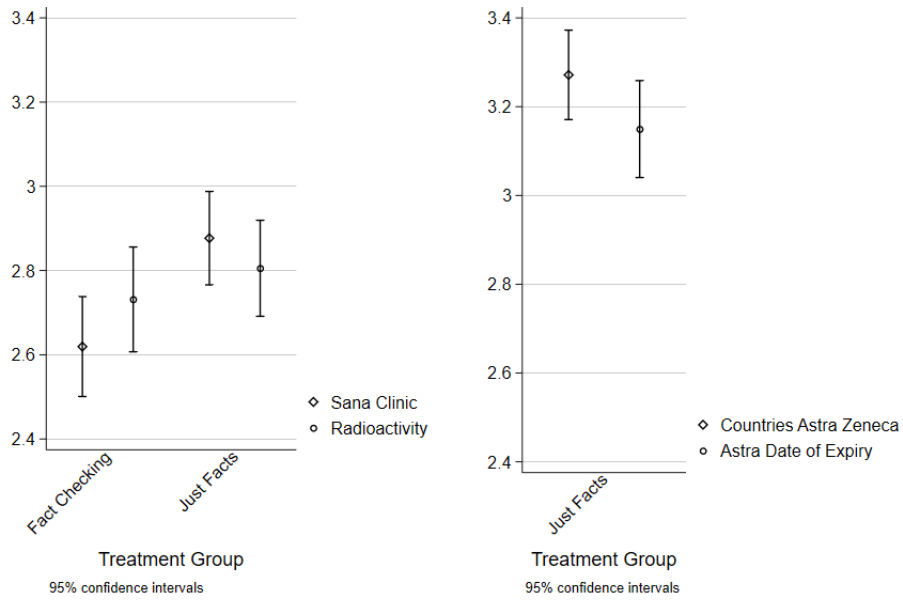


(a) Average credibility of fakes on Corona vaccines.

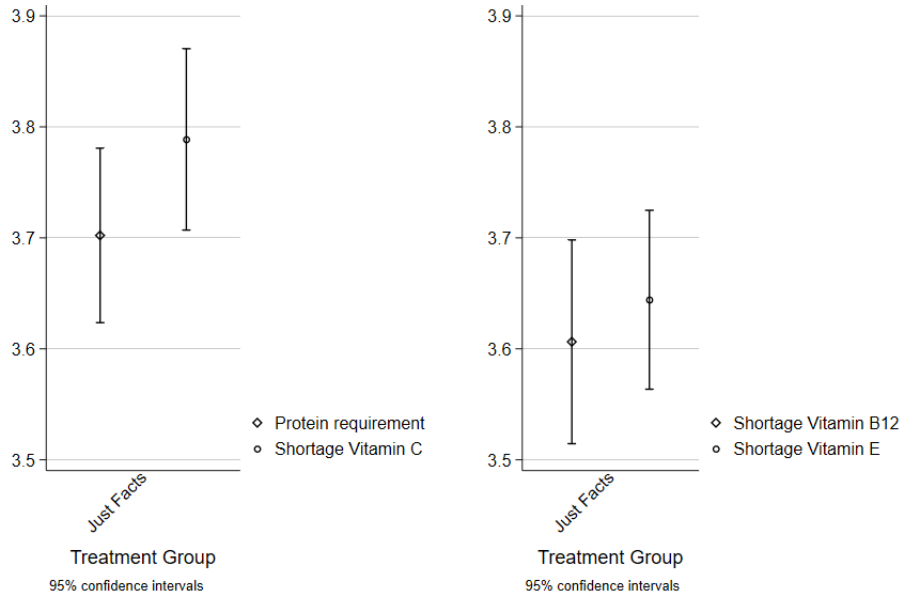


(b) Average credibility of fakes on nutrition.

Figure A.14: Average credibility of fakes on a 5-point Likert Scale.



(a) Average credibility of fact checks on Corona vaccines.



(b) Average credibility of fact checks on nutrition.

Figure A.15: Average credibility of fact checks on a 5-point Likert Scale.

D. Omitted tables

Table A.1: Balance table

| Variable | No Intervention | | Fact Checking | | Media Literacy | | Passive Control | | Just Facts | |
|----------------------------|-----------------|----------|---------------|----------|----------------|----------|-----------------|----------|------------|----------|
| | Mean | Std.Dev. | Mean | Std.Dev. | Mean | Std.Dev. | Mean | Std.Dev. | Mean | Std.Dev. |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| <i>General:</i> | | | | | | | | | | |
| Age | 40.741 | [11.833] | 40.542 | [11.983] | 40.165 | [11.906] | 40.241 | [12.510] | 41.075 | [12.114] |
| Male | 0.542 | [0.499] | 0.537 | [0.499] | 0.504 | [0.500] | 0.522 | [0.500] | 0.528 | [0.500] |
| Vaccinated | 0.788 | [0.409] | 0.835** | [0.371] | 0.825* | [0.380] | 0.830* | [0.376] | 0.814 | [0.389] |
| Supplement intake | 0.450 | [0.498] | 0.445 | [0.497] | 0.434 | [0.496] | 0.493 | [0.500] | 0.432 | [0.496] |
| <i>Prior knowledge:</i> | | | | | | | | | | |
| log dist. current events | 4.144 | [1.122] | 4.142 | [1.148] | 4.174 | [1.168] | 4.143 | [1.149] | 4.094 | [1.146] |
| log dist. health | -0.021 | [0.332] | -0.032 | [0.342] | -0.015 | [0.380] | -0.016 | [0.349] | -0.037 | [0.346] |
| log dist. nutrition | 0.542 | [0.537] | 0.515 | [0.535] | 0.540 | [0.536] | 0.555 | [0.544] | 0.536 | [0.531] |
| <i>Family status:</i> | | | | | | | | | | |
| Fam: Married | 0.411 | [0.492] | 0.376 | [0.485] | 0.383 | [0.487] | 0.403 | [0.491] | 0.405 | [0.491] |
| Fam: Common law marriage | 0.123 | [0.329] | 0.130 | [0.337] | 0.140 | [0.348] | 0.118 | [0.323] | 0.126 | [0.332] |
| Fam: Unmarried | 0.466 | [0.499] | 0.494 | [0.500] | 0.476 | [0.500] | 0.480 | [0.500] | 0.468 | [0.499] |
| <i>Household earnings:</i> | | | | | | | | | | |
| HH earnings < 1000 | 0.105 | [0.307] | 0.104 | [0.305] | 0.091 | [0.288] | 0.106 | [0.309] | 0.088 | [0.284] |
| HH earnings [1000,1999] | 0.222 | [0.416] | 0.189 | [0.392] | 0.207 | [0.406] | 0.224 | [0.417] | 0.261 | [0.439] |
| HH earnings [2000,2999] | 0.254 | [0.436] | 0.265 | [0.442] | 0.258 | [0.438] | 0.218 | [0.413] | 0.239 | [0.427] |
| HH earnings [3000,3999] | 0.162 | [0.369] | 0.170 | [0.376] | 0.170 | [0.376] | 0.167 | [0.373] | 0.201* | [0.401] |
| HH earnings > 4000 | 0.193 | [0.395] | 0.208 | [0.406] | 0.176 | [0.381] | 0.178 | [0.383] | 0.168 | [0.374] |
| HH earnings n.s. | 0.065 | [0.246] | 0.064 | [0.245] | 0.098** | [0.297] | 0.106*** | [0.309] | 0.043* | [0.203] |
| <i>Education:</i> | | | | | | | | | | |
| Education: no graduation | 0.002 | [0.040] | 0.010* | [0.099] | 0.007 | [0.081] | 0.005 | [0.070] | 0.002 | [0.041] |
| Education: CSE (cat 1) | 0.123 | [0.329] | 0.076*** | [0.265] | 0.088** | [0.284] | 0.118 | [0.323] | 0.100 | [0.300] |
| Education: CSE (cat2) | 0.325 | [0.469] | 0.315 | [0.465] | 0.334 | [0.472] | 0.337 | [0.473] | 0.352 | [0.478] |
| Education: high school | 0.286 | [0.452] | 0.292 | [0.455] | 0.302 | [0.459] | 0.273 | [0.446] | 0.281 | [0.450] |
| Education: college | 0.264 | [0.441] | 0.308* | [0.462] | 0.269 | [0.444] | 0.267 | [0.443] | 0.266 | [0.442] |
| <i>Personality traits:</i> | | | | | | | | | | |
| Big 5: conscientiousness | 5.440 | [1.086] | 5.340 | [1.118] | 5.377 | [1.111] | 5.349 | [1.118] | 5.401 | [1.114] |
| Big 5: extroversion | 4.360 | [1.331] | 4.443 | [1.345] | 4.375 | [1.344] | 4.378 | [1.306] | 4.386 | [1.294] |
| Big 5: tolerance | 5.059 | [1.073] | 4.938** | [1.087] | 5.097 | [1.088] | 5.050 | [1.101] | 5.035 | [1.073] |
| Big 5: openness | 4.563 | [1.210] | 4.462 | [1.260] | 4.528 | [1.213] | 4.556 | [1.312] | 4.539 | [1.213] |
| Big 5: neuroticism | 3.965 | [1.350] | 3.962 | [1.322] | 3.934 | [1.257] | 3.989 | [1.340] | 3.882 | [1.309] |
| <i>Party preferences:</i> | | | | | | | | | | |
| Vote: AfD | 0.095 | [0.294] | 0.092 | [0.290] | 0.093 | [0.291] | 0.085 | [0.279] | 0.128* | [0.334] |
| Vote: CDU/CSU | 0.173 | [0.379] | 0.157 | [0.364] | 0.153 | [0.361] | 0.164 | [0.370] | 0.128** | [0.334] |
| Vote: FDP | 0.105 | [0.307] | 0.096 | [0.294] | 0.114 | [0.318] | 0.108 | [0.311] | 0.136* | [0.343] |
| Vote: Greens | 0.176 | [0.381] | 0.199 | [0.400] | 0.176 | [0.381] | 0.195 | [0.396] | 0.173 | [0.378] |
| Vote: Left | 0.063 | [0.243] | 0.082 | [0.275] | 0.073 | [0.261] | 0.064 | [0.245] | 0.063 | [0.243] |
| Vote: SPD | 0.181 | [0.386] | 0.189 | [0.392] | 0.179 | [0.384] | 0.170 | [0.376] | 0.211 | [0.408] |
| Vote: Other | 0.206 | [0.404] | 0.185 | [0.388] | 0.210 | [0.408] | 0.214 | [0.411] | 0.161** | [0.368] |
| <i>State of residence:</i> | | | | | | | | | | |
| State: Baden-Württ. | 0.104 | [0.305] | 0.092 | [0.290] | 0.111 | [0.314] | 0.100 | [0.300] | 0.100 | [0.300] |
| State: Bayern | 0.146 | [0.353] | 0.145 | [0.352] | 0.148 | [0.356] | 0.162 | [0.369] | 0.143 | [0.350] |
| State: Berlin | 0.066 | [0.249] | 0.054 | [0.227] | 0.065 | [0.247] | 0.065 | [0.248] | 0.063 | [0.243] |
| State: Brandenburg | 0.026 | [0.159] | 0.028 | [0.165] | 0.036 | [0.186] | 0.029 | [0.169] | 0.033 | [0.179] |
| State: Bremen | 0.008 | [0.090] | 0.012 | [0.107] | 0.013 | [0.114] | 0.010 | [0.099] | 0.007 | [0.081] |
| State: Hamburg | 0.040 | [0.197] | 0.035 | [0.183] | 0.029 | [0.169] | 0.025 | [0.155] | 0.038 | [0.192] |
| State: Hessen | 0.084 | [0.278] | 0.091 | [0.287] | 0.067 | [0.250] | 0.088 | [0.284] | 0.078 | [0.269] |
| State: Mecklenburg-Vorp. | 0.023 | [0.149] | 0.016 | [0.127] | 0.013 | [0.114] | 0.011 | [0.107] | 0.022 | [0.145] |
| State: Niedersachsen | 0.087 | [0.283] | 0.112 | [0.316] | 0.095 | [0.293] | 0.077 | [0.267] | 0.081 | [0.274] |
| State: Nordrhein-Westf. | 0.204 | [0.403] | 0.224 | [0.417] | 0.192 | [0.395] | 0.221 | [0.415] | 0.211 | [0.408] |
| State: Rheinland-Pfalz | 0.052 | [0.222] | 0.058 | [0.233] | 0.060 | [0.238] | 0.043 | [0.202] | 0.033 | [0.179] |
| State: Saarland | 0.013 | [0.113] | 0.005 | [0.070] | 0.016 | [0.127] | 0.008 | [0.090] | 0.013 | [0.115] |
| State: Sachsen | 0.065 | [0.246] | 0.053 | [0.224] | 0.057 | [0.232] | 0.069 | [0.253] | 0.071 | [0.258] |
| State: Sachsen-Anhalt | 0.031 | [0.173] | 0.025 | [0.155] | 0.021 | [0.144] | 0.025 | [0.155] | 0.028 | [0.166] |
| State: Schleswig-Holstein | 0.034 | [0.181] | 0.025 | [0.155] | 0.044 | [0.205] | 0.049 | [0.216] | 0.050 | [0.218] |
| State: Thüringen | 0.018 | [0.132] | 0.026 | [0.160] | 0.031 | [0.173] | 0.018 | [0.133] | 0.028 | [0.166] |
| <i>N</i> | 618 | | 607 | | 613 | | 611 | | 602 | |

Notes: Table A.1 displays the mean values and standard deviations of all our control variables for each treatment group. We also conducted *t*-tests on the difference in means between the NoINTERVENTION and each of the other treatment groups respectively: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Table A.2: Summary statistics of our dependent variables

| Variable | Mean | Std. Dev. | Min. | Max. | N |
|--|--------|-----------|--------|-------|------|
| <i>Credibility:</i> | | | | | |
| Prob. credible/very credible – Corona vaccine fakes I | 0.238 | 0.426 | 0 | 1 | 1834 |
| Prob. credible/very credible – Nutrition fakes I | 0.876 | 0.329 | 0 | 1 | 1836 |
| Prob. credible/very credible – Corona vaccine fakes II | 0.382 | 0.486 | 0 | 1 | 1533 |
| Prob. credible/very credible – Nutrition fakes II | 0.798 | 0.401 | 0 | 1 | 1533 |
| Prob. credible/very credible – Corona vaccine facts I | 0.861 | 0.346 | 0 | 1 | 2438 |
| Prob. credible/very credible – Nutrition facts I | 0.946 | 0.226 | 0 | 1 | 2439 |
| Prob. credible/very credible – Corona vaccine facts II | 0.867 | 0.339 | 0 | 1 | 2006 |
| Prob. credible/very credible – Nutrition facts II | 0.956 | 0.206 | 0 | 1 | 2008 |
| <i>Factual knowledge:</i> | | | | | |
| Distance to truth in SD – Corona vaccine fakes I | -0.002 | 0.802 | -0.843 | 2.178 | 3051 |
| Distance to truth in SD – Nutrition fakes I | 0.001 | 0.727 | -1.014 | 2.894 | 3051 |
| Distance to truth in SD – Corona vaccine fakes II | -0.001 | 0.816 | -0.907 | 1.876 | 2525 |
| Distance to truth in SD – Nutrition fakes II | -0.004 | 0.74 | -1.048 | 2.894 | 2525 |
| Distance to truth in SD – Corona vaccine facts I | 0.002 | 0.803 | -0.676 | 2.965 | 3051 |
| Distance to truth in SD – Nutrition facts I | 0.002 | 0.776 | -0.567 | 2.991 | 3051 |
| Distance to truth in SD – Corona vaccine facts II | -0.001 | 0.734 | -0.563 | 2.718 | 2525 |
| Distance to truth in SD – Nutrition facts II | 0.001 | 0.808 | -0.793 | 1.789 | 2525 |
| <i>Attitudes:</i> | | | | | |
| Prob. likely/very likely – Covid vaccination I | 0.813 | 0.39 | 0 | 1 | 3051 |
| Prob. unlikely/very unlikely – Dietary supplements I | 0.548 | 0.498 | 0 | 1 | 3051 |
| Prob. likely/very likely – Covid vaccination II | 0.781 | 0.413 | 0 | 1 | 2525 |
| Prob. unlikely/very unlikely – Dietary supplements II | 0.553 | 0.497 | 0 | 1 | 2525 |
| <i>Likes and dislikes:</i> | | | | | |
| Corona vaccine fakes dislikes I | 3.656 | 3.627 | 0 | 21 | 1838 |
| Nutrition fakes dislikes I | 1.466 | 1.888 | 0 | 13 | 1838 |
| Corona vaccine fakes dislikes II | 4.855 | 4.574 | 0 | 26 | 1546 |
| Nutrition fakes dislikes II | 1.898 | 2.467 | 0 | 16 | 1546 |
| Corona vaccine facts likes I | 2.985 | 3.291 | 0 | 21 | 1838 |
| Nutrition facts likes I | 3.386 | 3.377 | 0 | 19 | 2440 |
| Corona vaccine facts likes II | 2.448 | 2.428 | 0 | 14 | 2028 |
| Nutrition facts likes II | 3.214 | 3.076 | 0 | 20 | 2028 |
| Corona vaccine fakes likes I | 0.671 | 1.419 | 0 | 17 | 1838 |
| Nutrition fakes likes I | 1.711 | 1.958 | 0 | 12 | 1838 |
| Corona vaccine fakes likes II | 1.42 | 1.998 | 0 | 15 | 1546 |
| Nutrition fakes likes II | 1.25 | 1.87 | 0 | 15 | 1546 |
| Corona vaccine facts dislikes I | 1.039 | 1.376 | 0 | 11 | 1838 |
| Corona vaccine facts dislikes II | 0.707 | 0.965 | 0 | 7 | 2028 |
| Nutrition facts dislikes I | 0.74 | 1.356 | 0 | 10 | 2440 |
| Nutrition facts dislikes II | 0.716 | 1.424 | 0 | 11 | 2028 |

Table A.3: Credibility of facts

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|----------------------|---------|-----------|---------|-----------------------|---------|-----------|---------|
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | -0.012 | -0.016 | -0.003 | -0.005 | -0.009 | -0.005 | -0.021 | -0.023 |
| | [0.020] | [0.020] | [0.013] | [0.013] | [0.022] | [0.022] | [0.014] | [0.014] |
| p-value | (0.526) | (0.427) | (0.841) | (0.684) | (0.687) | (0.815) | (0.121) | (0.092) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,224 | 1,224 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 |
| Panel B: Media literacy | | | | | | | | |
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | -0.006 | -0.010 | 0.008 | 0.005 | 0.014 | 0.016 | 0.002 | 0.001 |
| | [0.019] | [0.020] | [0.013] | [0.013] | [0.021] | [0.021] | [0.012] | [0.012] |
| p-value | (0.759) | (0.606) | (0.546) | (0.720) | (0.510) | (0.436) | (0.861) | (0.937) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,018 | 1,018 | 1,020 | 1,020 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 0.869 | 0.869 | 0.943 | 0.943 | 0.866 | 0.866 | 0.961 | 0.961 |
| Std.Dev. DV | 0.338 | 0.338 | 0.231 | 0.231 | 0.341 | 0.341 | 0.194 | 0.194 |

Notes: Table A.3 presents the OLS coefficients of comparing the NOINTERVENTION to the FACTCHECKING (Panel A) and to the MEDIALITERACY group (Panel B), respectively. The outcome is a dummy variable equal to one if participant *i* perceives the **facts** on Corona vaccines and nutrition in Wave I and Wave II of the survey as *Credible* or *Very credible* on average. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, income, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.4: Credibility – truth discernment

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking * fake | -0.068 | -0.068 | -0.013 | -0.013 | -0.000 | -0.000 | -0.005 | -0.005 |
| | [0.032] | [0.032] | [0.019] | [0.019] | [0.035] | [0.036] | [0.026] | [0.026] |
| p-value | (0.034) | (0.035) | (0.504) | (0.501) | (0.994) | (0.994) | (0.851) | (0.852) |
| Fact checking | -0.012 | -0.009 | -0.003 | -0.001 | -0.009 | 0.006 | -0.021 | -0.018 |
| | [0.020] | [0.020] | [0.013] | [0.013] | [0.022] | [0.023] | [0.014] | [0.014] |
| p-value | (0.526) | (0.639) | (0.842) | (0.910) | (0.687) | (0.801) | (0.121) | (0.190) |
| fake | -0.570 | -0.570 | -0.042 | -0.042 | -0.463 | -0.464 | -0.130 | -0.130 |
| | [0.023] | [0.023] | [0.013] | [0.013] | [0.024] | [0.025] | [0.017] | [0.017] |
| p-value | (0.000) | (0.000) | (0.001) | (0.001) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,224 | 1,224 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 |
| Panel B: Media literacy | | | | | | | | |
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy * fake | -0.098 | -0.098 | -0.067 | -0.067 | -0.066 | -0.066 | -0.074 | -0.074 |
| | [0.031] | [0.031] | [0.020] | [0.020] | [0.035] | [0.035] | [0.026] | [0.027] |
| p-value | (0.002) | (0.002) | (0.001) | (0.001) | (0.057) | (0.057) | (0.005) | (0.005) |
| Media literacy | -0.006 | -0.009 | 0.008 | 0.006 | 0.014 | 0.020 | 0.002 | 0.003 |
| | [0.019] | [0.020] | [0.013] | [0.013] | [0.021] | [0.022] | [0.012] | [0.012] |
| p-value | (0.759) | (0.663) | (0.546) | (0.660) | (0.510) | (0.347) | (0.861) | (0.790) |
| fake | -0.570 | -0.570 | -0.042 | -0.042 | -0.463 | -0.464 | -0.130 | -0.130 |
| | [0.023] | [0.023] | [0.013] | [0.013] | [0.025] | [0.025] | [0.017] | [0.017] |
| p-value | (0.000) | (0.000) | (0.001) | (0.001) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,018 | 1,018 | 1,020 | 1,020 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 0.584 | 0.584 | 0.922 | 0.922 | 0.635 | 0.635 | 0.895 | 0.895 |
| Std.Dev. DV | 0.493 | 0.493 | 0.268 | 0.268 | 0.482 | 0.482 | 0.305 | 0.305 |

Notes: All estimates are OLS estimates. The dependent variable is a dummy equal to one if participant i considered the fakes or facts on Corona vaccines or nutrition in Wave I or Wave II of the survey on average as *Very credible* or *Credible*. Panel A shows the estimates from comparing the FACTCHECKING, and Panel B from comparing the MEDIALITERACY to the NOINTERVENTION group, respectively. Robust standard errors in squared parentheses, p-values in round parentheses. Standard errors are clustered on the participant level; N corresponds to the number of clusters. Control variables include age, gender, family status, income, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.5: Distance to truth on facts

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | 0.059 | 0.059 | -0.015 | -0.005 | 0.056 | 0.062 | -0.003 | 0.010 |
| | [0.045] | [0.044] | [0.042] | [0.042] | [0.041] | [0.041] | [0.045] | [0.045] |
| p-value | (0.193) | (0.178) | (0.730) | (0.902) | (0.171) | (0.124) | (0.951) | (0.819) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,225 | 1,225 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 |
| Panel B: Media literacy | | | | | | | | |
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | 0.009 | 0.002 | 0.021 | 0.022 | 0.032 | 0.043 | -0.005 | 0.013 |
| | [0.044] | [0.042] | [0.042] | [0.040] | [0.043] | [0.042] | [0.045] | [0.044] |
| p-value | (0.831) | (0.971) | (0.613) | (0.581) | (0.447) | (0.305) | (0.908) | (0.761) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,020 | 1,020 | 1,020 | 1,020 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | -0.094 | -0.094 | -0.100 | -0.100 | -0.180 | -0.180 | -0.206 | -0.206 |
| Std.Dev. DV | 0.805 | 0.805 | 0.753 | 0.753 | 0.648 | 0.648 | 0.726 | 0.726 |

Notes: Table A.5 presents OLS estimates for participants' distance to truth on topics that the **facts** on Corona vaccines and nutrition in Wave I and Wave II of the survey are dealing with. Panel A shows the estimates from comparing the FACTCHECKING, and Panel B from comparing the MEDIALITERACY to the NOINTERVENTION group, respectively. The dependent variable is equal to participant *i*'s mean average standardized distance to the correct answer. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.6: Distance to truth – truth discernment

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking * fake | -0.381 | -0.381 | -0.066 | -0.066 | -0.104 | -0.104 | 0.023 | 0.023 |
| | [0.071] | [0.071] | [0.049] | [0.049] | [0.075] | [0.075] | [0.059] | [0.060] |
| p-value | (0.000) | (0.000) | (0.181) | (0.181) | (0.164) | (0.167) | (0.698) | (0.700) |
| Fact checking | 0.059 | 0.060 | -0.0015 | -0.010 | 0.056 | 0.058 | -0.003 | 0.009 |
| | [0.045] | [0.044] | [0.042] | [0.042] | [0.041] | [0.041] | [0.045] | [0.044] |
| p-value | (0.193) | (0.168) | (0.730) | (0.815) | (0.171) | (0.157) | (0.951) | (0.842) |
| fake | 0.492 | 0.492 | 0.232 | 0.232 | 0.514 | 0.514 | 0.349 | 0.349 |
| | [0.052] | [0.052] | [0.034] | [0.034] | [0.053] | [0.053] | [0.043] | [0.043] |
| p-value | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,225 | 1,225 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 |
| Panel B: Media literacy | | | | | | | | |
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy * fake | -0.229 | -0.229 | -0.100 | -0.100 | -0.147 | -0.147 | -0.021 | -0.021 |
| | [0.071] | [0.071] | [0.050] | [0.050] | [0.075] | [0.076] | [0.059] | [0.060] |
| p-value | (0.001) | (0.001) | (0.045) | (0.046) | (0.052) | (0.053) | (0.721) | (0.723) |
| Media literacy | 0.003 | 0.005 | 0.021 | 0.021 | 0.032 | 0.032 | -0.005 | -0.003 |
| | [0.044] | [0.043] | [0.042] | [0.040] | [0.043] | [0.043] | [0.045] | [0.044] |
| p-value | (0.831) | (0.900) | (0.613) | (0.608) | (0.447) | (0.448) | (0.908) | (0.944) |
| fake | 0.492 | 0.492 | 0.232 | 0.232 | 0.514 | 0.514 | 0.349 | 0.349 |
| | [0.052] | [0.052] | [0.034] | [0.034] | [0.053] | [0.053] | [0.042] | [0.043] |
| p-value | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,020 | 1,020 | 1,020 | 1,020 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 0.152 | 0.152 | 0.016 | 0.016 | 0.077 | 0.077 | -0.031 | -0.031 |
| Std.Dev. DV | 0.859 | 0.859 | 0.705 | 0.705 | 0.795 | 0.795 | 0.678 | 0.678 |

Notes: All estimates are OLS estimates. The dependent variable is the average standardized distance to the true value that the fakes or facts on Corona vaccines or nutrition in Wave I or Wave II of the survey are covering. Panel A shows the estimates from comparing the FACTCHECKING, and Panel B from comparing the MEDIALITERACY to the NOINTERVENTION group, respectively. Robust standard errors in squared parentheses, p-values in round parentheses. Standard errors are clustered on the participant level; *N* corresponds to the number of clusters. Control variables include age, gender, family status, income, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.7: Heterogeneity in baseline beliefs on nutrition – Fact checking

| Panel A: Fact checking – AfD supporters | | | | | | |
|--|---------------|---------|---------|----------------|---------|---------|
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Suppl. | Cred. | Dist. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Fact checking | -0.019 | -0.054 | 0.091 | -0.029 | -0.107 | 0.037 |
| | [0.067] | [0.156] | [0.099] | [0.075] | [0.108] | [0.117] |
| p-value | (0.783) | (0.728) | (0.361) | (0.704) | (0.329) | (0.755) |
| Controls | yes | yes | yes | yes | yes | yes |
| <i>N</i> | 114 | 115 | 115 | 99 | 99 | 99 |
| Baseline: No Intervention | | | | | | |
| Mean DV | 0.881 | 0.221 | 0.576 | 0.885 | 0.273 | 0.596 |
| Std.Dev. DV | 0.326 | 0.663 | 0.498 | 0.323 | 0.614 | 0.495 |
| Panel B: Fact checking – non-AfD supporters | | | | | | |
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Suppl. | Cred. | Dist. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Fact checking | -0.009 | -0.073 | -0.003 | -0.017 | 0.060 | 0.023 |
| | [0.018] | [0.036] | [0.030] | [0.026] | [0.041] | [0.033] |
| p-value | (0.631) | (0.043) | (0.907) | (0.523) | (0.144) | (0.491) |
| Controls | yes | yes | yes | yes | yes | yes |
| <i>N</i> | 1,109 | 1,110 | 1,110 | 923 | 923 | 923 |
| Baseline: No Intervention | | | | | | |
| Mean DV | 0.903 | 0.122 | 0.555 | 0.825 | 0.129 | 0.549 |
| Std.Dev. DV | 0.296 | 0.629 | 0.497 | 0.380 | 0.572 | 0.498 |

Notes: Table A.7 displays the effect heterogeneity between AfD supporters (Panel A) and non-AfD supporters (Panel B) for our **Fact checking** intervention. In columns 1 and 4, the dependent variable is a dummy equal to one if participant i perceives the fakes on nutrition as *Very credible* or *Credible* on average. In columns 2 and 5, the dependent variable is equal to participant i 's mean average standardized distance to the correct answer. In columns 3 and 6, the dependent variable is a dummy equal to one if participant i states to be *Unlikely* or *Very unlikely* to consume dietary supplements in the near future. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.8: Heterogeneity in baseline beliefs on nutrition – Media literacy

| Panel A: Media literacy – AfD supporters | | | | | | |
|---|---------------|---------|---------|----------------|---------|---------|
| | Wave I | | | Wave II | | |
| | Cred. | Dist. | Suppl. | Cred. | Dist. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Media literacy | 0.043 | 0.044 | 0.006 | -0.012 | -0.156 | 0.001 |
| | [0.063] | [0.152] | [0.101] | [0.068] | [0.153] | [0.100] |
| p-value | (0.495) | (0.774) | (0.952) | (0.866) | (0.309) | (0.990) |
| Controls | yes | yes | yes | yes | yes | yes |
| <i>N</i> | 116 | 116 | 116 | 101 | 101 | 101 |
| Baseline: No Intervention | | | | | | |
| Mean DV | 0.881 | 0.221 | 0.576 | 0.885 | 0.273 | 0.596 |
| Std.Dev. DV | 0.326 | 0.663 | 0.498 | 0.323 | 0.614 | 0.495 |
| Panel B: Media literacy – non-AfD supporters | | | | | | |
| | Wave I | | | Wave II | | |
| | Cred. | Dist. | Suppl. | Cred. | Dist. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Media literacy | -0.067 | -0.083 | -0.033 | -0.069 | -0.027 | -0.003 |
| | [0.020] | [0.037] | [0.030] | [0.026] | [0.038] | [0.033] |
| p-value | (0.001) | (0.026) | (0.268) | (0.009) | (0.481) | (0.937) |
| Controls | yes | yes | yes | yes | yes | yes |
| Mean DV | 0.869 | 0.079 | 0.537 | 0.788 | 0.122 | 0.542 |
| Std.Dev. DV | 0.337 | 0.632 | 0.499 | 0.409 | 0.605 | 0.499 |
| <i>N</i> | 1,115 | 1,115 | 1,115 | 919 | 919 | 919 |
| Baseline: No Intervention | | | | | | |
| Mean DV | 0.903 | 0.122 | 0.555 | 0.825 | 0.129 | 0.549 |
| Std.Dev. DV | 0.296 | 0.629 | 0.497 | 0.380 | 0.572 | 0.498 |

Notes: Table A.8 displays the effect heterogeneity between AfD supporters (Panel A) and non-AfD supporters (Panel B) for our **Media literacy** intervention. In columns 1 and 4, the dependent variable is a dummy equal to one if participant i perceives the fakes on and nutrition as *Very credible* or *Credible* on average. In columns 2 and 5, the dependent variable is equal to participant i 's mean average standardized distance to the correct answer. In columns 3 and 6, the dependent variable is a dummy equal to one if participant i states to be *Unlikely* or *Very unlikely* to consume dietary supplements in the near future. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.9: Likes of fakes

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | 0.035 | 0.041 | -0.060 | -0.046 | -0.214 | -0.186 | 0.001 | 0.034 |
| | [0.082] | [0.083] | [0.108] | [0.109] | [0.120] | [0.123] | [0.117] | [0.119] |
| p-value | (0.668) | (0.617) | (0.578) | (0.671) | (0.076) | (0.131) | (0.992) | (0.773) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,225 | 1,225 | 1,225 | 1,225 | 1,030 | 1,030 | 1,030 | 1,030 |
| Panel B: Media literacy | | | | | | | | |
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | -0.006 | -0.031 | 0.128 | 0.098 | -0.113 | -0.128 | 0.032 | 0.018 |
| | [0.082] | [0.082] | [0.114] | [0.113] | [0.131] | [0.134] | [0.113] | [0.111] |
| p-value | (0.941) | (0.706) | (0.264) | (0.389) | (0.388) | (0.339) | (0.776) | (0.873) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,026 | 1,026 | 1,026 | 1,026 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 0.662 | 0.662 | 1.688 | 1.688 | 1.529 | 1.529 | 1.239 | 1.239 |
| Std.Dev. DV | 1.468 | 1.468 | 1.918 | 1.918 | 2.058 | 2.058 | 1.751 | 1.751 |

Notes: Table A.9 compares the absolute number **likes on fakes** on Corona vaccines and dietary supplements for participants from the NOINTERVENTION to the FACTCHECKING (Panel A) and the MEDIALITERACY group (Panel B) in Wave I and Wave II of the survey, respectively. All estimates are OLS estimates. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.10: Dislikes of facts

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | -0.156 | -0.146 | -0.126 | -0.140 | 0.027 | 0.026 | -0.065 | -0.081 |
| | [0.079] | [0.078] | [0.070] | [0.070] | [0.059] | [0.060] | [0.081] | [0.078] |
| p-value | (0.048) | (0.064) | (0.072) | (0.045) | (0.647) | (0.660) | (0.423) | (0.301) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,225 | 1,225 | 1,225 | 1,225 | 1,030 | 1,030 | 1,030 | 1,030 |
| Panel B: Media literacy | | | | | | | | |
| | <u>Wave I</u> | | | | <u>Wave II</u> | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | -0.025 | -0.041 | 0.322 | 0.327 | -0.043 | -0.055 | 0.235 | 0.226 |
| | [0.081] | [0.082] | [0.086] | [0.085] | [0.059] | [0.061] | [0.094] | [0.093] |
| p-value | (0.755) | (0.621) | (0.000) | (0.000) | (0.467) | (0.366) | (0.013) | (0.016) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,231 | 1,231 | 1,231 | 1,231 | 1,026 | 1,026 | 1,026 | 1,026 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 1.099 | 1.099 | 0.696 | 0.696 | 0.704 | 0.704 | 0.665 | 0.665 |
| Std.Dev. DV | 1.491 | 1.491 | 1.315 | 1.315 | 0.915 | 0.915 | 1.366 | 1.366 |

Notes: Table A.10 compares the absolute number of **dislikes on facts** on Corona vaccines and dietary supplements for participants from the NOINTERVENTION to the FACTCHECKING (Panel A) and the MEDIALITERACY group (Panel B) in Wave I and Wave II of the survey, respectively. All estimates are OLS estimates. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.11: Distance to truth on topics covered by fakes – Comparison to PASSIVECONTROL

| Panel A: Fact checking | | | | | | | | |
|----------------------------------|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Fact checking | 0.512 | 0.533 | 0.022 | 0.053 | 0.757 | 0.764 | 0.240 | 0.259 |
| | [0.041] | [0.040] | [0.043] | [0.042] | [0.043] | [0.043] | [0.046] | [0.047] |
| p-value | (0.000) | (0.000) | (0.615) | (0.210) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,218 | 1,218 | 1,218 | 1,218 | 1,030 | 1,030 | 1,030 | 1,030 |
| Panel B: Media literacy | | | | | | | | |
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Media literacy | 0.614 | 0.604 | 0.024 | 0.039 | 0.691 | 0.676 | 0.194 | 0.198 |
| | [0.041] | [0.040] | [0.044] | [0.044] | [0.042] | [0.041] | [0.046] | [0.046] |
| p-value | (0.000) | (0.000) | (0.593) | (0.373) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,224 | 1,224 | 1,224 | 1,224 | 1,028 | 1,028 | 1,028 | 1,028 |
| Panel C: No Intervention | | | | | | | | |
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| No Intervention | 0.834 | 0.840 | 0.102 | 0.109 | 0.805 | 0.808 | 0.220 | 0.222 |
| | [0.041] | [0.041] | [0.043] | [0.043] | [0.041] | [0.042] | [0.045] | [0.045] |
| p-value | (0.000) | (0.000) | (0.020) | (0.012) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,229 | 1,229 | 1,229 | 1,229 | 1,026 | 1,026 | 1,026 | 1,026 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | -0.436 | -0.436 | 0.030 | 0.030 | -0.471 | -0.471 | -0.077 | -0.077 |
| Std.Dev. DV | 0.589 | 0.589 | 0.879 | 0.879 | 0.419 | 0.419 | 0.833 | 0.833 |

Notes: Table A.11 presents OLS estimates for participants' distance to truth on topics that the **fakes** on Corona vaccines and nutrition in Wave I and Wave II of the survey are dealing with. Panel A shows the estimates from comparing the FACTCHECKING, and Panel B from comparing the MEDIALITERACY to the PASSIVECONTROL group, respectively. The dependent variable is equal to participant *i*'s mean average standardized distance to the correct answer. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.12: Attitudes towards Corona vaccination and the intake of dietary supplements – Comparison to PASSIVECONTROL

| Panel A: Fact checking | | | | | | | | | | |
|----------------------------------|--------------------|---------|---------|-------------|---------|--------------------|---------|---------|-------------|---------|
| | Wave I | | | | | Wave II | | | | |
| | Corona vaccination | | | Supplements | | Corona vaccination | | | Supplements | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| Fact checking | -0.006 | -0.014 | -0.011 | -0.034 | -0.038 | 0.012 | 0.007 | 0.007 | -0.003 | 0.003 |
| | [0.021] | [0.020] | [0.018] | [0.028] | [0.028] | [0.025] | [0.024] | [0.021] | [0.031] | [0.031] |
| p-value | (0.778) | (0.464) | (0.541) | (0.231) | (0.173) | (0.637) | (0.784) | (0.747) | (0.913) | (0.929) |
| Controls | no | yes | yes + | no | yes | no | yes | yes + | no | yes |
| <i>N</i> | 1,218 | 1,218 | 1,218 | 1,218 | 1,218 | 1,030 | 1,030 | 1,030 | 1,030 | 1,030 |
| Panel B: Media literacy | | | | | | | | | | |
| | Wave I | | | | | Wave II | | | | |
| | Corona vaccination | | | Supplements | | Corona vaccination | | | Supplements | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| Media literacy | 0.001 | -0.000 | 0.001 | -0.074 | -0.075 | 0.019 | 0.023 | 0.015 | -0.036 | -0.033 |
| | [0.021] | [0.021] | [0.019] | [0.028] | [0.028] | [0.025] | [0.024] | [0.022] | [0.031] | [0.031] |
| p-value | (0.980) | (0.981) | (0.972) | (0.009) | (0.007) | (0.451) | (0.338) | (0.493) | (0.242) | (0.282) |
| Controls | no | yes | yes + | no | yes | no | yes | yes + | no | yes |
| <i>N</i> | 1,224 | 1,224 | 1,224 | 1,224 | 1,224 | 1,028 | 1,028 | 1,028 | 1,028 | 1,028 |
| Panel C: No Intervention | | | | | | | | | | |
| | Wave I | | | | | Wave II | | | | |
| | Corona vaccination | | | Supplements | | Corona vaccination | | | Supplements | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) |
| No Intervention | -0.053 | -0.051 | -0.032 | -0.033 | -0.035 | -0.045 | -0.038 | -0.028 | -0.015 | -0.016 |
| | [0.022] | [0.021] | [0.019] | [0.028] | [0.028] | [0.027] | [0.025] | [0.022] | [0.031] | [0.031] |
| p-value | (0.018) | (0.015) | (0.092) | (0.249) | (0.210) | (0.093) | (0.122) | (0.219) | (0.637) | (0.614) |
| Controls | no | yes | yes + | no | yes | no | yes | yes + | no | yes |
| <i>N</i> | 1,229 | 1,229 | 1,229 | 1,229 | 1,229 | 1,026 | 1,026 | 1,026 | 1,026 | 1,026 |
| Baseline: No Intervention | | | | | | | | | | |
| Mean DV | 0.836 | 0.836 | 0.836 | 0.589 | 0.589 | 0.785 | 0.785 | 0.785 | 0.569 | 0.569 |
| Std.Dev. DV | 0.370 | 0.370 | 0.370 | 0.492 | 0.492 | 0.411 | 0.411 | 0.411 | 0.496 | 0.496 |

Notes: Table A.12 presents the OLS estimates of comparing the PASSIVECONTROL to the FACTCHECKING (Panel A) and to the MEDIALITERACY group (Panel B), respectively. The dependent variable is a dummy equal to one if participant i states to be *Likely* or *Very likely* to get vaccinated or boosted against Covid-19, or *Unlikely* or *Very unlikely* to consume dietary supplements in the near future. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition. In columns 3 and 8 (“yes +”), we also control for participants’ Corona vaccination status.

Table A.13: Distance to truth on topics covered by fakes – JUSTFACTS group

| Panel A: Comparison to NoIntervention group | | | | | | | | |
|--|---------------|---------|-----------|---------|----------------|---------|-----------|---------|
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Just facts | -0.631 | -0.641 | -0.596 | -0.588 | -0.731 | -0.729 | -0.540 | -0.535 |
| | [0.042] | [0.042] | [0.061] | [0.061] | [0.046] | [0.047] | [0.046] | [0.047] |
| p-value | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| Mean DV | 0.398 | 0.398 | 0.212 | 0.212 | 0.333 | 0.333 | 0.143 | 0.143 |
| Std.Dev. DV | 0.841 | 0.841 | 0.769 | 0.769 | 0.844 | 0.844 | 0.577 | 0.577 |
| <i>N</i> | 1,220 | 1,220 | 1,220 | 1,220 | 984 | 984 | 984 | 984 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | 0.398 | 0.398 | 0.212 | 0.212 | 0.333 | 0.333 | 0.143 | 0.143 |
| Std.Dev. DV | 0.841 | 0.841 | 0.769 | 0.769 | 0.844 | 0.844 | 0.577 | 0.577 |
| Panel B: Comparison to PassiveControl group | | | | | | | | |
| | Wave I | | | | Wave II | | | |
| | Corona | | Nutrition | | Corona | | Nutrition | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Just facts | 0.202 | 0.222 | -0.277 | -0.239 | 0.074 | 0.080 | -0.319 | -0.302 |
| | [0.034] | [0.033] | [0.069] | [0.067] | [0.033] | [0.034] | [0.053] | [0.055] |
| p-value | (0.000) | (0.000) | (0.000) | (0.000) | (0.024) | (0.019) | (0.000) | (0.000) |
| Controls | no | yes | no | yes | no | yes | no | yes |
| <i>N</i> | 1,213 | 1,213 | 1,213 | 1,213 | 992 | 992 | 992 | 992 |
| Baseline: No Intervention | | | | | | | | |
| Mean DV | -0.436 | -0.436 | 0.030 | 0.030 | -0.471 | -0.471 | -0.077 | -0.077 |
| Std.Dev. DV | 0.589 | 0.589 | 0.879 | 0.879 | 0.419 | 0.419 | 0.833 | 0.833 |

Notes: Table A.13 compares distance to truth on topics that the Corona vaccine and nutrition **fakes** are dealing with between participants from the NOINTERVENTION (Panel A) and the PASSIVECONTROL (Panel B) and the JUSTFACTS group, respectively. All estimates are OLS estimates. The dependent variable is equal to participant i 's mean average standardized distance to the correct answer. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.14: List experiments – Sample split

| Panel A: Vaccination | | | | |
|-----------------------------|---------------------------------|---------|-----------------------------------|---------|
| | Likely to get vaccinated | | Unlikely to get vaccinated | |
| | Wave I | Wave II | Wave I | Wave II |
| | (1) | (2) | (3) | (4) |
| List experiment | 7.06 | 0.00 | 52.71 | 35.69 |
| N | 2,480 | 1,973 | 571 | 552 |

| Panel B: Supplements | | | | |
|-----------------------------|-----------------------------------|---------|-------------------------------------|---------|
| | Likely to take supplements | | Unlikely to take supplements | |
| | Wave I | Wave II | Wave I | Wave II |
| | (1) | (2) | (3) | (4) |
| List experiment | 79.23 | 79.01 | 19.23 | 12.99 |
| N | 918 | 780 | 2,133 | 1,745 |

Notes: Panel A splits participants who directly report to be *Very likely* or *Likely* to get vaccinated against Covid-19 in the main experiment from those who did not and displays the respective indirectly elicited proportions from the list experiments for each of those subsamples. Analogously, Panel B splits participants who directly report to be *Very unlikely* or *Unlikely* to consume dietary supplements in the main experiment from those who did not and displays the respective proportions from the list experiments.

Table A.15: 2SLS estimates for our main specifications

| Panel A: Fact checking | | | | | | | | | | | | |
|---------------------------|---------|---------|---------|-----------|---------|---------|---------|---------|---------|-----------|---------|---------|
| | Wave I | | | | | | Wave II | | | | | |
| | Corona | | | Nutrition | | | Corona | | | Nutrition | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Suppl. | Cred. | Dist. | Vaccine | Cred. | Dist. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Second stage | | | | | | | | | | | | |
| \widehat{D}_i | -0.101 | -0.458 | 0.025 | -0.015 | -0.115 | 0.005 | 0.023 | -0.072 | 0.047 | -0.026 | 0.043 | 0.030 |
| | [0.034] | [0.070] | [0.026] | [0.025] | [0.049] | [0.040] | [0.042] | [0.077] | [0.031] | [0.034] | [0.054] | [0.044] |
| p-value | (0.003) | (0.000) | (0.332) | (0.558) | (0.020) | (0.907) | (0.578) | (0.344) | (0.122) | (0.453) | (0.420) | (0.497) |
| First stage | | | | | | | | | | | | |
| Fact checking | 0.699 | 0.696 | 0.698 | 0.699 | 0.697 | 0.697 | 0.704 | 0.704 | 0.705 | 0.704 | 0.704 | 0.704 |
| | [0.018] | [0.019] | [0.019] | [0.019] | [0.019] | [0.019] | [0.020] | [0.020] | [0.020] | [0.020] | [0.020] | [0.020] |
| p-value | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) |
| F-statistic | 1429.95 | 1409.5 | 1408.42 | 1421.1 | 1409.5 | 1409.5 | 1186.3 | 1186.3 | 1186.61 | 1186.3 | 1186.3 | 1186.3 |
| Controls | yes | yes | yes + | yes | yes | yes | yes | yes | yes + | yes | yes | yes |
| N | 1,221 | 1,225 | 1,225 | 1,223 | 1,225 | 1,225 | 1,022 | 1,022 | 1,022 | 1,022 | 1,022 | 1,022 |
| Panel B: Media literacy | | | | | | | | | | | | |
| | Wave I | | | | | | Wave II | | | | | |
| | Corona | | | Nutrition | | | Corona | | | Nutrition | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Suppl. | Cred. | Dist. | Vaccine | Cred. | Dist. | Suppl. |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| Second stage | | | | | | | | | | | | |
| \widehat{D}_i | -0.142 | -0.296 | 0.046 | -0.082 | -0.109 | -0.051 | -0.058 | -0.189 | 0.065 | -0.093 | -0.056 | -0.017 |
| | [0.031] | [0.065] | [0.025] | [0.025] | [0.048] | [0.038] | [0.039] | [0.071] | [0.030] | [0.033] | [0.050] | [0.042] |
| p-value | (0.000) | (0.000) | (0.071) | (0.001) | (0.024) | (0.181) | (0.135) | (0.008) | (0.030) | (0.005) | (0.262) | (0.690) |
| First stage | | | | | | | | | | | | |
| Media Literacy | 0.740 | 0.740 | 0.741 | 0.740 | 0.740 | 0.740 | 0.731 | 0.731 | 0.733 | 0.731 | 0.731 | 0.731 |
| | [0.017] | [0.017] | [0.017] | [0.017] | [0.017] | [0.017] | [0.019] | [0.019] | [0.019] | [0.019] | [0.019] | [0.019] |
| p-value | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) | (0.000) |
| F-statistic | 1799.59 | 1799.59 | 1797.85 | 1799.59 | 1799.59 | 1799.59 | 1411.51 | 1411.51 | 1412.8 | 1411.51 | 1411.51 | 1411.51 |
| Controls | yes | yes | yes + | yes | yes | yes | yes | yes | yes + | yes | yes | yes |
| N | 1,231 | 1,231 | 1,231 | 1,231 | 1,231 | 1,231 | 1,020 | 1,020 | 1,020 | 1,020 | 1,020 | 1,020 |
| Baseline: No Intervention | | | | | | | | | | | | |
| Mean DV | 0.299 | 0.398 | 0.783 | 0.901 | 0.132 | 0.557 | 0.403 | 0.334 | 0.741 | 0.831 | 0.143 | 0.554 |
| Std.Dev. DV | 0.458 | 0.841 | 0.412 | 0.299 | 0.632 | 0.497 | 0.491 | 0.844 | 0.439 | 0.375 | 0.577 | 0.498 |

Notes: Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits (“big 5”), political preferences, and prior knowledge on current events, health, and nutrition. In columns 3 and 9 (“yes +”), we also control for participants’ Corona vaccination status.

Table A.16: Heterogeneity in vaccination status – FACTCHECKING

| Panel A: Fact checking – Fully vaccinated | | | | | | |
|--|---------------|---------|---------|----------------|---------|---------|
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Vaccine |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Fact checking | -0.059 | -0.323 | 0.018 | 0.037 | -0.049 | 0.036 |
| | [0.026] | [0.054] | [0.018] | [0.032] | [0.060] | [0.023] |
| p-value | (0.023) | (0.000) | (0.310) | (0.243) | (0.414) | (0.119) |
| Controls | yes | yes | yes | yes | yes | yes |
| <i>N</i> | 991 | 994 | 994 | 835 | 835 | 835 |
| Panel B: Fact checking – Not fully vaccinated | | | | | | |
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Vaccine |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Fact checking | -0.056 | -0.299 | 0.034 | -0.051 | -0.043 | 0.049 |
| | [0.069] | [0.115] | [0.065] | [0.075] | [0.142] | [0.071] |
| p-value | (0.417) | (0.010) | (0.596) | (0.496) | (0.765) | (0.488) |
| Controls | yes | yes | yes | yes | yes | yes |
| <i>N</i> | 230 | 231 | 231 | 187 | 187 | 187 |
| Baseline: No Intervention | | | | | | |
| Mean DV | 0.299 | 0.398 | 0.783 | 0.403 | 0.334 | 0.741 |
| Std.Dev. DV | 0.458 | 0.841 | 0.412 | 0.491 | 0.844 | 0.439 |

Notes: Table A.16 displays the effect heterogeneity between fully vaccinated (Panel A) and not fully vaccinated (Panel B) participants for our **Fact checking** intervention. The NOINTERVENTION group is the omitted category in all specifications. In columns 1 and 4, the dependent variable is a dummy equal to one if participant i perceives the fakes on Corona vaccines as *Very credible* or *Credible* on average. In columns 2 and 5, the dependent variable is equal to participant i 's mean average standardized distance to the correct answer. In columns 3 and 6, the dependent variable is a dummy equal to one if participant i states to be *Likely* or *Very likely* to get vaccinated or boosted against Covid-19. All estimates are OLS estimates. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition.

Table A.17: Heterogeneity in vaccination status – MEDIA LITERACY

| Panel A: Media literacy – Fully vaccinated | | | | | | |
|---|---------------|---------|---------|----------------|---------|---------|
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Vaccine |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Media literacy | -0.093 | -0.211 | 0.011 | -0.020 | -0.187 | 0.041 |
| | [0.025] | [0.054] | [0.018] | [0.031] | [0.058] | [0.023] |
| p-value | (0.000) | (0.000) | (0.558) | (0.530) | (0.001) | (0.077) |
| Controls | yes | yes | yes | yes | yes | yes |
| <i>N</i> | 993 | 993 | 993 | 835 | 835 | 835 |
| Panel B: Media literacy – Not fully vaccinated | | | | | | |
| | <u>Wave I</u> | | | <u>Wave II</u> | | |
| | Cred. | Dist. | Vaccine | Cred. | Dist. | Vaccine |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Media literacy | -0.081 | -0.292 | 0.119 | -0.047 | 0.058 | 0.056 |
| | [0.066] | [0.106] | [0.061] | [0.073] | [0.125] | [0.069] |
| p-value | (0.216) | (0.006) | (0.054) | (0.526) | (0.642) | (0.423) |
| Controls | yes | yes | yes | yes | yes | yes |
| <i>N</i> | 238 | 238 | 238 | 185 | 185 | 185 |
| Baseline: No Intervention | | | | | | |
| Mean DV | 0.299 | 0.398 | 0.783 | 0.403 | 0.334 | 0.741 |
| Std.Dev. DV | 0.458 | 0.841 | 0.412 | 0.491 | 0.844 | 0.439 |

Notes: Table A.17 displays the effect heterogeneity between fully vaccinated (Panel A) and not fully vaccinated (Panel B) participants for our **media literacy** intervention. The NOINTERVENTION group is the omitted category in all specifications. In columns 1 and 4, the dependent variable is a dummy equal to one if participant i perceives the fakes on Corona vaccines as *Very credible* or *Credible* on average. In columns 2 and 5, the dependent variable is equal to participant i 's mean average standardized distance to the correct answer. In columns 3 and 6, the dependent variable is a dummy equal to one if participant i states to be *Likely* or *Very likely* to get vaccinated or boosted against Covid-19. All estimates are OLS estimates. Robust standard errors in squared parentheses, p-values in round parentheses. Control variables include age, gender, family status, household earnings, education, personality traits ("big 5"), political preferences, and prior knowledge on current events, health, and nutrition.