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Article

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Article

India's Total Natural Resource Rents (NRR) and GDP: An Augmented Autoregressive Distributed Lag (ARDL) Bound Test

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Abstract: Utilizing natural resources wisely, reducing pollution, and taking other environmental factors into account are now critical to the prospects for long-term economic growth and, by extension, sustainable development. We investigate the impact of total natural resource rents (NRR) on India's GDP in this study. The data sample consists of NRR and GDP data from the World Bank's official website collected between 1993 and 2020. In the study, the Granger causality test and an augmented autoregressive distributed lag (ARDL) bound test were used. The NRR have a significant impact on India's GDP, according to the results of the ARDL model on the framed time series data set. Furthermore, the ARDL bound test reveals that the NRR have a significant short-term and long-term impact on the GDP of the Indian economy. This research contributes to understanding whether an exclusive policy is required for effective management of the complex interactions between various forces in the economic, political, and social environments. This is significant because there is no standard policy in India to improve the efficiency of utility extraction from natural resources.

Keywords: total natural resource rents; GDP; sustainable development; ARDL bound test; India



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1. Introduction

Countries endowed with natural resources should fare better and develop faster than countries devoid of valuable natural capital. However, this is not always the case, as some countries with abundant natural resources experience higher levels of poverty, inequality, and civil unrest than others (Ampofo et al. 2020). Existing empirical studies show that natural resource abundance has a negative effect on economic growth, a phenomenon known in most countries as the resource curse (Sachs and Warner 1995). According to Auty and Furlonge (2019), abundant natural resources have not been the panacea for igniting or sustaining economic growth and development in many countries. The purely economic view of the natural resource curse is based on explaining why resource-rich countries grow more slowly than resource-poor countries (Frankel 2010; Sachs and Warner 1997).

The economic value derived from natural resources such as land, minerals, and water is referred to as resource rent. The difference between the market price of the resource and the cost of production can be used to calculate resource rent (OECD 2022). The economic value of resource rent is not accounted for in the calculation of a country's GDP. While resource rent helps a country's economy grow, it does not directly contribute to GDP. For example, if a country extracts oil from its land, the revenue from selling the oil increases the country's GDP, but the value of the oil itself does not. Resource rent, on the other hand, is the difference between the market price of oil and the cost of extracting it (OECD 2022).

In recent years, after the introduction of the Sustainable Development Goals (SDG) in 2015, several governments have prioritized environmental challenges such as air pollution, poor sanitation, and significant damage to natural resource rents (OECD 2011) and forest reserves (Pereira-Romeiro et al. 2021). Governments are also focusing on green finance

to solve environmental issues (Bhatnagar et al. 2022b). Human health and economic well-being are jeopardized by poor environmental conditions (Tallis et al. 2008). These aspects, including well-being, environmental and physical capital, and approaches to water, food, and land, are all susceptible to climate change (Ibarrarán et al. 2009). These environmental concerns have generated a global movement to combat climate change (Han and Ahn 2020). Nonetheless, the majority of the Belt and Road Initiative countries must struggle to update their manufacturing progress (Tekdal 2018), resulting in significant fossil fuel energy consumption in the production division (York 2012), which contributes to global warming. Carbon (CO₂) emissions have historically been employed as a proxy for environmental quality in the present ecological sustainability literature (Alola and Kirikkaleli 2019). However, numerous researchers have questioned this measure, claiming that CO₂ emissions only account for a small percentage of the whole environment and do not adequately capture environmental contamination (Han et al. 2019).

GDP has been the focus of all policymakers because it represents the welfare of the nation. Furthermore, many researchers have contributed to the discovery of the factors responsible for driving economic growth (Hegwood and Papell 2007; Auty and Furlonge 2019; Fleissig and Strauss 1999; Zaidi and Rupeika-Apoga 2021). The aim of this study is to investigate the impact of total natural resource rents (NRR) on India's GDP.

We analyzed NRR and GDP data from the World Bank's official website collected between 1993 and 2020 to achieve the study's aim. To identify causality between NRR and GDP, we used the Granger causality test. The augmented autoregressive distributed lag (ARDL) bound test was used to estimate the effect of NRR on GDP. We discovered that NRR has a significant short-term and long-term impact on the Indian economy's GDP.

This study helps to determine whether an exclusive policy is necessary for the effective management of the complex interactions between various forces in the economic, political, and social environments. This study makes a contribution by emphasizing the need for India to develop standard policies to improve utility extraction efficiency from natural resources.

The article is organized as follows: Section 2 includes a review of the literature, which we used to develop the hypotheses, and Section 3 discusses the research methodology used. The research findings are presented in Section 4. The scientific and practical implications of the study's findings are discussed in Section 5. Section 6 summarizes the results and makes recommendations for additional study.

2. Literature Review and Hypotheses Development

Natural resource rents are calculated as the difference between an item's value and the average cost of producing it (Hamilton and Clemens 1999). This is accomplished by calculating the value of units of specific supplies and subtracting valuations of typical unit expenses for extraction or harvesting costs (Hung 1993). The rents on all of these modules are then multiplied by the amount of material extracted or harvested by nations to obtain the rents on each item as a percentage of GDP (Hamilton and Ruta 2006).

Records for the influence of natural resources on financial production is vital in establishing a systematic outline for sustainable progress (Shi and Gill 2005). Profits from natural assets, particularly minerals and fossil fuels, account for a significant portion of GDP in some countries, and the majority of these revenues are in the form of monetary rents—income above the cost of extracting the resources (Mehtar et al. 2018). As natural resources are not created, they generate economic rents (Mehtar et al. 2018). For created services and commodities, competitive forces increase supply until revenues are pushed to zero, whereas natural assets that have a predetermined supply typically command revenues far in excess of their production costs (Arrow and Debreu 1954). Rents from non-renewable resources such as fossil fuels and minerals, as well as rents from overharvesting of the jungle, demonstrate a "liquidation of a country's capital stock" (Mehtar et al. 2018). When governments use such rents to fund current expenditure rather than new wealth

expenditure to replenish what has been depleted, they borrow from their future prospects (Schumpeter and Keynes 1936).

According to a review of the literature, natural resource rent has had a significant impact on major economic indicators (Dramani et al. 2022; Ben-Salha et al. 2021; Frankel 2010; Ampofo et al. 2020). Many studies of the variables, i.e., natural resources and economic indicators (Bhatnagar and Pathak 2021; Özen and Taneja 2022), are conducted around the world (Cloutens and Kirat 2017; Hegwood and Papell 2007; Ibrahim 2017; Fleissig and Strauss 1999; Sachs and Warner 1995; Stijns 2005), but significantly less research has been conducted in India (Ampofo et al. 2020). This paper attempts to fill a gap in the assessment of the trend analysis of natural resource rent and the measurement of its impact on India's economic growth.

Revenue collected through taxation systems or schemes is a significant economic source for a country's project development and the achievement of its macroeconomic goals. As a result, natural resources are important in both established and emerging economies. Rent-based resources are crucial in determining economic well-being (Mpofu 2021).

The consequences of financial globalization on environmental dilapidation (CO₂ emissions) were examined for the G7 nations from 1996 to 2017. The empirical evidence indicates that economic globalization, financial growth, and the exploitation of natural assets contribute to the rise in carbon emissions. For the long-term and short-term findings of the empirical research, this work employs innovative econometric methodologies such as CS-ARDL (Wang et al. 2020). This study aimed to determine if NRR may amplify the positive relationship between industrial growth and human capital with a sample of 14 Sub-Saharan African nations, covering the period of 1995 to 2015. The empirical data specify that NRR has a negative and statistically negligible effect on industrial improvement. It was also disclosed that the industrial sector has yet to reap significant benefits despite the enormous rents generated by natural resources (Bolarinwa et al. 2019).

Decentralization of fiscal policy and rent from natural resources affect CO₂ emissions. The article of Tufail et al. (2021) accomplishes this purpose by analyzing panel data from seven highly decentralized members of the OECD nations from 1990 to 2018. At the 1st difference, all variables are shown to be stationary. Long-run results indicate that devolution and natural capital rents help the environment reduce CO₂ emissions (Tufail et al. 2021).

Given China's severe environmental degradation, the current study establishes the different factors' dependency trends, emphasizing the intersection of eco-revolution and biological impact throughout many quantiles of the distributions, which has not been investigated previously in the research. As experimental evidence demonstrates, growing private-sector debt and ecological benefits contribute to environmental deterioration (Afshan and Yaqoob 2022).

Total resource rents (coal rentals, mineral rents, natural gas rental costs, and forest rents) have been evaluated in a worldwide sample of 90 countries. Economic expansion, FDI inflows, investment, government expenditures, and demographic considerations were used as control variables. The global sample consisted of 90 economies. Natural rents are influenced by economic complexity, but only in low- and lower-middle income economies and high-income economies, and not in upper-middle income economies, according to the findings (Canh et al. 2020).

Development in underdeveloped nations is directly linked to the amount of natural resource rents they obtain from their natural resources. Research on the consequences of natural resource rents shows that this connection varies by area, income level, degree of development, and export structure. When CO₂ emissions are added as an interaction variable for Africa, natural resource rents have a detrimental influence on inclusive personal expansion (Nchofoung et al. 2021).

It is crucial for ecological progress to focus on both sustainable production and consumption (Rupeika-Apoga and Petrovska 2022). Panel data analysis yielded robust findings by combining different econometric methodologies. Granger causality arises between two variables, particularly financing and natural resource rents. The second place indicates

that a deepening of the financial markets has a favorable impact on natural resource rents (Nguyen and Nguyen 2020).

Researchers have long been interested in what happens when natural resources are abundant and how this affects economic development, both theoretically and empirically. However, opinions differ on whether a resource surplus helps or hinders economic growth. The studies found no discernible pattern despite using a variety of samples, data sources, and approaches. Torvik (2002) develops a novel and relatively simple mechanism to explain how an abundance of natural resources may be detrimental to people's well-being. Gyamfi et al. (2022) agree with the assertion that the majority of profits from natural resource rents go toward increasing production, which has a negative impact on the environment and a country's economic development. A country's stable monetary system supports economic growth, which in turn supports the growth of other industries (Rupeika-Apoga et al. 2022). The empirical findings have important policy implications, particularly for the need to promote globalization and human capital development in order to ensure the appropriate management of natural resource rent, which is critical for financial sector growth (Guan et al. 2020).

While most empirical results of the resource effect include how the development path is decided through the channel of institutions, the question of why resource rents frequently fail to inspire more robust governance is more important than ever for the nation's economic development (Fagbemi and Adeoye 2020).

The relationship between NRR and GDP is complicated and is affected by a number of factors, including the type of resource, the country's economic policies, and its ability to diversify its economy (Dramani et al. 2022; Ben-Salha et al. 2021; Ibrahim 2017). The Granger causality test is frequently used to investigate the relationship between NRR and GDP, as this statistical hypothesis test determines whether NRR is useful in forecasting GDP. The Granger causality test had an insignificant impact in Australia, Brazil, Canada, China, Saudi Arabia, the United States, and Venezuela (Ampofo et al. 2020). Long-run Granger causality tests conducted by Mohamed (2020) revealed a unidirectional causal relationship running from resource rents to GDP growth, as well as development expenditure to GDP growth. The Toda–Yamamoto test method also revealed a unidirectional causality connecting oil resource rent to Ghana's economic growth (Adabor et al. 2022). Considering this, we proposed the following hypotheses:

H1a. *Total natural resource rents (NRR) do not Granger-cause GDP in India.*

H1b. *GDP does not Granger-cause total natural resource rents (NRR) in India.*

Natural resource rents, according to one study (Meher et al. 2018), have a significant impact on a country's GDP and overall economic activity. Natural resources, according to Stijns (2005), are not significant determinants of economic growth due to the coexistence of positive and negative channels of effect running from natural resources to factors affecting economic growth. According to Ben-Salha et al. (2021), there has been a massive body of research on the impact of natural capital on GDP development over the last several decades. The research has focused on developed and emerging countries, particularly those with abundant natural resources (Rudra and Jensen 2011). In empirical studies of the natural resources–economic growth nexus, two primary indicators, resource dependency and resource abundance, have been considered (Xu et al. 2022). The former describes a scenario in which a country's earnings are primarily derived from the extraction and sale of natural resources, whereas the latter describes a scenario in which such resources are abundant (Mehlum et al. 2006). The yearly per capita rent of resource output captures the abundance of resources. Natural resource rents over GDP, the percentage of total natural resource exports in GDP, and the share of natural resources in total exports, on the other hand, are indicators of resource dependency (Yilanci et al. 2021). According to the study by Topcu et al. (2020), the majority of natural resource rent earnings are invested in expanding production, which has a negative impact on the environment and slows a country's economic growth. A country's stable monetary system promotes economic

development, which in turn encourages the expansion of other businesses (Friedman 2017). Natural resource rent management is critical to financial sector expansion, and Ncube et al.'s (2014) empirical findings have important policy implications on this front. Even though most empirical data on resource impact conclude that the development path is chosen through institutions (Khan et al. 2020), it is more important than ever for a country's economic growth to address why resource rents frequently fail to inspire more strong governance (Collier 2005; Kaur et al. 2021). Therefore, we hypothesize:

H2. Total natural resources rents (NRR) have no significant impact on India's GDP.

3. Materials and Methods

For meeting the research objective on the basis of the research gap drawn from the conducted literature review, data of total natural resources rents and India's GDP was obtained from the official website of World Bank and time series data is exhibited in Figure 1 (World Bank 2022).

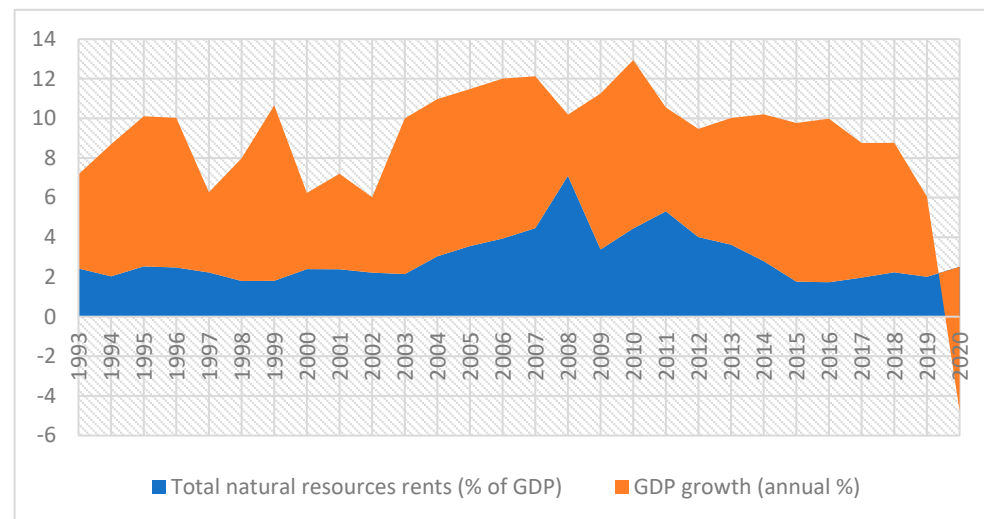


Figure 1. Time series data of total natural resource rents and India's GDP from 1993 to 2020. Source: (World Bank 2022).

Due to a lack of the literature on causality between GDP and NRR in the Indian economy, the Granger causality test was used to distinguish which variable was acting as a dependent and which variable was acting as an independent. The autoregressive distributed lag (ARDL) model was used to examine the short- and long-run effects of NRR on India's GDP. The ARDL model was helpful for forecasting and for separating short-run dynamics from long-run relationships. When lag weights fell to zero relatively quickly, the regressor was not highly autocorrelated, and the sample was long in comparison to the length of the lag distribution. The model was best suited to estimating dynamic relationships (Bhatnagar et al. 2022a). Data collection from 1993 to 2020 was completed in order to run the analysis. The following regression equation served as the research's foundation:

$$y_t = c_0 + \sum_{k=1}^p \beta_k y_{t-k} + \sum_{j=0}^l \alpha_{j+1} x_{t-j} + u_t \tag{1}$$

where,

y_t = India's GDP of Current Year,

y_{t-k} = India's GDP of Previous Year,

x_{t-j} = Total natural resources rents of Previous Year,

u_t = Error Terms.

The above equation will serve as the foundation for short-term assessment using the ARDL bound test and long-run analysis using F-bound. Before the application of final model of ARDL, in the initial analysis, unit root test was applied by employing the following equation:

$$\Delta y_t = \alpha + \beta t + \gamma y_{t-1} + \delta_1 \Delta y_{t-1} + \dots + \delta_{p-1} \Delta y_{t-p+1} + \varepsilon_t \tag{2}$$

Once unit root test projected results at level and first difference of regressor and regressand, granger causality test was levied in following way:

$$Y_t = \sum_{i=1}^n \alpha_i Y_{t-i} + \sum_{j=1}^n \beta_j X_{t-j} + u_{1t} \tag{3}$$

$$X_t = \sum_{i=1}^n \lambda_i Y_{t-i} + \sum_{j=1}^n \sigma_j X_{t-j} + u_{2t} \tag{4}$$

By the application of these mathematical models an attempt to meet the research objectives has been made in upcoming sections of research.

4. Results

Table 1 displays descriptive statistics.

Table 1. Descriptive Statistics.

Particulars	Total Natural Resources Rents (% of GDP)	GDP Growth (Annual %)
Mean	2.937767584	5.993638435
Standard Error	0.236934487	0.58438374
Median	2.44921419	6.727153743
Standard Deviation	1.253739457	3.092268092
Sample Variance	1.571862625	9.56212195
Kurtosis	3.308835486	12.4088011
Skewness	1.702171742	−3.071150803
Range	5.37124876	16.09751034
Minimum	1.729645283	−7.251754782
Maximum	7.100894043	8.845755561
Sum	82.25749234	167.8218762
Count	28	28
Largest (1)	7.100894043	8.845755561
Smallest (1)	1.729645283	−7.251754782
Confidence Level (95.0%)	0.48614941	1.199056391

Table 2 shows the unit root test for the dependent variable in the study. Table 2 applies unit root test on GDP and NRR by using following equation:

$$\Delta y_t = \alpha + \beta t + \gamma y_{t-1} + \delta_1 \Delta y_{t-1} + \dots + \delta_{p-1} \Delta y_{t-p+1} + \varepsilon_t \tag{5}$$

After performing the unit root test on India’s GDP, the likelihood is greater than 0.05, which is marginally related to data stationary. Furthermore, the Dickey–Fuller test t-statistic is less than two, implying that we cannot rule out the possibility that GDP does not have a unit root. The unit root measurement has a lag length of zero, which is automatically determined from SIC with a maximum lag of six. Table 2 shows the statistics for using the unit root test with the least squares method on the dependent variable, GDP. These calculations were performed at a level that yielded a log-likelihood of −67.75. The values in Table 2 show that the series is not stationary at the level of India’s GDP. Table 2 also includes the augmented Dickey–Fuller test statistics for the first difference with a leg length of zero and a maximum lag of six. It demonstrates critical test values at 1%, 5%, and 10%

significance levels. The probability value is 0.0058, which is less than 0.05, and this is why the results are statistically significant. The supporting calculations regarding the augmented Dickey–Fuller test provide insight into the further analysis of whether the series chosen can be considered to have concluded. Log-likelihood and the augmented Dickey–Fuller test were used to validate the unit root test statistics calculations. The negative log-likelihood value (−66.46) is a good sign, and the F statistic (15.55) indicates a favorable Dickey–Fuller test result. The p-value of the unit root test, on the other hand, is 0.0006, which is less than 0.05. Based on this value, the series of India’s GDP can be interpreted as stationary at first difference.

Table 2. Unit root tests.

		At Level	
		GDP Growth (Annual %)	Total Natural Resources Rents (% of GDP)
With Constant	t-Statistic	−1.471	−7.0906
	Prob.	0.5327	0 ***
With Constant and Trend	t-Statistic	−1.3441	−7.0344
	Prob.	0.8543	0 ***
		At First Difference	
		GDP growth (annual %)	Total natural resources rents (% of GDP)
With Constant	t-Statistic	−3.9438	−2.2933
	Prob.	0.0058 ***	0.1811
With Constant and Trend	t-Statistic	−4.1889	−2.2241
	Prob.	0.0144 **	0.4583

Notes: (**) significant at the 5%; (***) significant at the 1%; and (no) not significant, lag length based on SIC, probability based on MacKinnon (1996) one-sided p-values.

The calculations regarding the augmented Dickey–Fuller to verify the presence of the unit root in the series of NRR are constructed in Table 2. The calculation is initiated by assuming that NRR have a unit root test. The calculations were performed with a lag length of zero and a max lag of six. The probability value was less than 0.05, and the critical test values of the independent variable were tested at three levels of significance: 1%, 5%, and 10%. Based on the values given in Table 2, the sum of the square of residuals (25.39), the log-likelihood (−36.58) and the value of F statistics (50.28) indicate a good sign for the time-series data collected as an independent variable. Moreover, the p-value shown in Table 2 for the NRR is almost zero, less than 0.05. This value of probability results in the rejection of the null hypothesis for the NRR, which shows the series is stationary at a level. As a result, as an interpretation, the independent variable of the research is stationary at the level and the dependent variable is stationary at the first difference. Therefore, the ARDL model is an optimal model to measure the impact of an independent variable on a dependent variable.

Table 3 displays the Granger causality test results. The purpose of this test was to determine which variables are dependent and which are independent. The test revealed that unidirectional causality exists from NRR to GDP.

Table 3. Pairwise Granger causality.

Hypotheses	Observations	F-Value	Prob.
NRR does not Granger-Cause GDP	27	9.59353	0.0049
GDP does not Granger-Cause NRR		0.73753	0.3989

It means that NRR have an effect on GDP. As a result, the upcoming analysis will assess the long and short run effects of NNR on India’s GDP.

Table 4 includes statistical values for the F-bound test, which assumes no levels of relationships as the null hypothesis. In Table 4, the value of the F statistic is analogized with a lower and upper bound at a 10%, 5%, 2.5%, and 1% level of significance for testing the significance of the continuation of a relationship between the regressor and controlled variable. On the test application, the F value is 8.629. This is greater than all of the upper bound values, which are 3.8, 4.6, 5.39, and 6.44 at 10%, 5%, 2.5%, and 1%, respectively. Hence, based on these calculations, it can be deduced that there is cointegration between the variables chosen for research.

Table 4. F-Bound Test (Null Hypothesis: No levels relationship).

Test Statistic	Value	Signif.	I(0)	I(1)
Asymptotic: n = 1000				
F-statistic	8.629709	10%	3.8	3.8
		5%	4.6	4.6
k	1	2.5%	5.39	5.39
		1%	6.44	6.44
Finite Sample: n = 35				
Actual Sample Size	27	10%	3.98	3.98
		5%	4.945	4.945
		1%	7.35	7.35
Finite Sample: n = 30				
		10%	4.025	4.025
		5%	5.07	5.07
		1%	7.595	7.595

The application of the autoregressive distributed lag model on India’s GDP as a dependent variable and NRR as an independent variable shows specific statistics exhibited in Table 5. The test includes a total of 27 observations, which are taken into account after adjustments. Four lags are selected automatically, and the model selection criteria are conducted based on the Akaike info criterion (AIC). The t statistics and probability values show that the model is a good fit, with t-values greater than 2 and p-values less than 0.05.

Table 5. ARDL Bound Test (Short Run).

Variable	Coefficient	Std. Error	t-Statistic	Prob.
NRR	0.653355	0.151156	4.32238	0.0002
C (Constant)	1.026983	0.484875	2.118035	0.0443
R-squared	0.827695			
Adjusted R-squared	0.804802			
S.E. of regression	0.982458			
Sum squared residuals	24.13059			
Log likelihood	−36.7945			
F-statistic	18.68296			
Prob(F-statistic)	0.000216			
Mean dependent var	2.956891			
S.D. dependent var	1.273454			
Akaike info criterion	2.873669			
Schwarz criterion	2.969656			
Hannan-Quinn criter.	2.902211			
Durbin-Watson stat	2.274002			

Notes: p-values and any subsequent tests do not account for model selection.

Tables 6 and 7 contain the conditional error correction regression of ARDL long-run form and bound test. This regression test contains GDP as the dependent variable and NRR as the independent variable. The ARDL model is run as a case of restricted constant

and no trend, which includes time-series data of 27 observations each of dependent and independent variables. Table 6 shows a compatible t-Bound distribution because the p-values of both constant and NRR are less than 0.05, indicating that the results are significant in the model.

Table 6. Conditional Error Correction Regression.

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C (Constant)	1.026983	0.484875	2.118035	0.0443
NRR	−0.346645	0.151156	−2.293286	0.0305

Table 7. Restricted Constant and No Trend.

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C (Constant)	2.962638	0.545454	5.431508	0

Furthermore, the level equation’s restricted constant and no trend value in Table 7 show that the relationship is significant because the probability is less than 0.05. The coefficient value of 2.962638 indicates that the relationship is also positive. Hence, the values shown in Table 7 reveal a direct and significant relationship between dependent and independent variables in the long run.

Table 8 contains the test application of ARDL on the error correction model. The statistical calculations shown in the table determine whether or not there is a long-term relationship between the variables. The coefficient of the cointegration equation, which has the value −0.3466 in Table 8, is used to estimate the probability of predicting a long relationship. When this value is combined with the probability value of 0.0276, it can be deduced that the long-term adjustment is negative and significant, indicating that the model will adjust monotonically.

Table 8. ARDL Error Correction Regression.

ECM Regression				
Included Observations: 27				
Variable	Coefficient	Std. Error	t-Statistic	Prob.
NRR	1.432795	0.521007	2.750048	0.0114
CointEq(−1) *	−0.34665	0.148217	−2.33876	0.0276
R-squared	0.173804			
Adjusted R-squared	0.173804			
S.E. of regression	0.963379			
Sum squared residuals	24.13059			
Log likelihood	−36.7945			
Durbin-Watson stat	2.274002			
Mean dependent var	0.003049			
S.D. dependent var	1.059878			
Akaike info criterion	2.799594			
Schwarz criterion	2.847588			
Hannan-Quinn criter.	2.813866			

* The coefficient of the cointegration equation.

Table 9 displays the F-bound test values, which are computed under the assumption of the null hypothesis, which states that no levels of relationships exist between the variables. Furthermore, the statistics value is 11.62971, which is greater than the upper bound values at a 10%, 5%, 2.5%, and 1% level of significance, indicating long-run cointegration among the variables.

Table 9. F-bound Test for Long-Run Analysis.

F-Bound Test		Null Hypothesis: No Levels Relationship		
Test Statistic	Value	Signif.	I(0)	I(1)
F-statistic	11.62971	10%	3.8	3.8
k	0	5%	4.6	4.6
		2.5%	5.39	5.39
		1%	6.44	6.44

5. Discussion

Natural resource rent is an important component of the economy because it provides governments and people with a consistent and reliable source of income. Natural resource rent is generated by the extraction of natural resources such as oil, gas, minerals, and timber, and it can be a significant source of government revenue in resource-rich countries (Dramani et al. 2022). The term “Dutch disease”, which is frequently used in the “resource curse” literature, was coined to describe the Dutch economic experience, in which the manufacturing sector declined and general inflation increased as a result of the booming natural gas sector (Horsley et al. 2015). Identifying the impact of natural resource rents on GDP in India is difficult due to a lack of reliable data and research. Some findings indicate that natural resource rents have boosted India’s economic growth (Mehtar et al. 2018). While natural resource rent is a curse in India, Australia, and the Democratic Republic of the Congo, it has boosted growth in Brazil and Canada, according to Ampofo et al. (2020). Furthermore, according to Dramani et al. (2022), natural resource rent can have a double threshold effect on economic growth. When aggregate natural resource rent exceeds 6% of GDP, the negative impact on economic growth is significant; however, as rents rise above 6% to around 15% of GDP, the negative impact on economic growth decreases significantly.

In our study, we first established a causal relationship between GDP and NRR to investigate the role of NRR in the growth of India’s GDP. We discovered a statistically significant causal relationship between NRR and GDP, as the *p*-value of the test is less than 0.05. This means that NRR can be used to forecast India’s GDP. As a result, we were able to reject H1a, which states that total natural resource rents (NRR) do not Granger-cause GDP, implying that NRR does.

Our findings are consistent with Mohamed’s (2020) long-run Granger causality tests, which revealed a unidirectional causal relationship running from resource rents to GDP growth as well as development expenditure to GDP growth. The Toda–Yamamoto test method also revealed a unidirectional causality running from oil resource rent to Ghana’s economic growth (Adabor et al. 2022).

We have not found evidence for GDP Granger-causing NRR. As a result, we failed to reject H1b, that GDP does not Granger-cause total natural resource rents (NRR) in India. When investigating a series of unit root and causality tests to detect causality between GDP and energy consumption in Turkey using Hsiao’s version of the Granger causality method for the 1950–2000 period, Altinay and Karagol (2004) discovered that there was no evidence of causality between GDP and energy consumption in Turkey based on detrended data. According to Indra de Soysa (2015), resource dependence measured by rents per GDP and resource abundance measured by rents per capita are not causally related.

Based on the results of the application of autoregressive distributed lag to the framed time series data set, it is possible to conclude that NRR have a significant impact on India’s GDP. The final result values of the ARDL bound tests in Table 5 show that NRR have a significant short-term impact on the Indian economy’s GDP. The *p*-value is 0.0002 and is less than 0.05, which is used to interpret the significant short-term influence of NRR on GDP. Table 9 also shows the F-bound test results under the null hypothesis, that there are no correlations between the variables. The statistical value of 11.62971 is greater than the upper limits of significance at 10%, 5%, 2.5%, and 1%, indicating that the variables are cointegrated in the long run. Cointegration indicates that the NRR and GDP series are linked and move in tandem over time; when one series rises or falls, the other follows suit. Cointegration can

be used to forecast NRR and GDP values in the future. This means we can reject H2, which claims that NRR have no effect on India's GDP. Our findings are supported by the [Adabor et al. \(2022\)](#) study, the main finding of which revealed that a 1% increase in oil resource rent generates a 0.84% increase in Ghana's economic growth in the long run.

[Ampofo et al. \(2020\)](#) examined the impact of total natural resource rent on economic growth using the nonlinear autoregressive distributed lag (NARDL) and nonlinear Granger causality in the case of the world's top ten mineral-rich countries from 1981 to 2017. They discovered nonlinear cointegration between total natural resource rent, trade openness, and economic growth in Australia, Brazil, Canada, the Democratic Republic of the Congo (DRC), India, and Saudi Arabia. These findings support our findings of cointegration between total natural resource rent and economic growth. Furthermore, they discovered that total natural resource rent has a negative impact on economic growth in Australia, the Democratic Republic of the Congo, and India, confirming the resource curse phenomenon in these countries. This highlights the implications of our study findings, which suggest that Indian policymakers should consider this when developing natural resource development policies. Moreover, [Cloutens and Kirat \(2017\)](#) discovered that the effect of natural resource rent on economic growth is dependent on the economic performance of different groups, demonstrating the resource curse in low-income countries.

GDP growth has been the focus of policymakers in every country, developed or developing. Long-run cointegration exists between the variables of Indian resource rent and GDP in the current study. However, when other countries are considered, different results can be observed for the United States, China, and Australia ([Ben-Salha et al. 2021](#)). In addition, a study conducted in Nigeria discovered that the impact of natural resource rent on the country's economic development was minimal. The study also suggested bringing strong political stability to make natural resource rent more impactful ([Fagbemi and Adeoye 2020](#)).

Our findings have important policy implications, particularly the need for public policies and regulations focusing on resource sustainability and the prevention of over-exploitation in order to ensure the appropriate use of natural resources to boost economic growth. In order to combat environmental degradation and promote sustainability, India is implementing a new national statistical framework that measures the services provided by natural ecosystems beginning in 2021 ([UN 2021](#)). This reflects a global shift away from treating natural resources as if they were free and limitless and toward measuring and valuing nature's contribution to our economies.

As a result of our findings, it is possible to advocate for the use of natural resource rents to invest in clean energy for a more sustainable environment while maintaining a focused revenue goal.

6. Conclusions

The gross domestic product (GDP) is the most commonly used measure of economic performance, but it says nothing about the wealth and assets that support this income. Mineral resource extraction, for example, boosts GDP while depleting a country's wealth. Utilizing natural resources wisely, reducing pollution, and considering other environmental factors are now critical to the prospects for long-term economic growth and, by extension, sustainable development.

Our findings allowed us to reject H1a, which states that total natural resource rents (NRR) do not Granger-cause GDP, implying that NRR does. However, we failed to reject H1b, which states that GDP does not Granger-cause total natural resource rents (NRR) in India. Furthermore, the H2 was rejected, indicating that the NRR has a significant impact on the GDP of India.

Total natural resource rents are an indicator that demonstrates the efficiency of extracting energy from natural resources at the lowest possible cost. If a country wishes to ensure macroeconomic development, all activities and policies articulated by the government must be profitable. The higher the rent value, the more profitable the use of

natural resources. The advancement of technology used to extract utility from natural resources can only increase the value of this variable. The current study demonstrates that total natural resource rents have a significant impact on India's GDP, highlighting the importance of policy formulation. In India, there is no exclusive policy regarding total natural resource rents. As a result, policymakers in the country must prioritize NRR in order to provide a significant boost to India's GDP. Natural resources are classified into two types: nonrenewable and renewable. The government must encourage the use of renewable energy for sustainable development, but this concept can only be expanded if it is economically profitable. Furthermore, if renewable energy is primarily contributing to GDP, the government will be interested in developing renewable energy policies.

The study's limitation is that there are numerous factors that influence a nation's GDP, and the current study only considered the variable of natural resource rent as an independent variable, but there are other related variables that are not included in the current research. Further research can be directed toward understanding what role NRR play in green and digital transformation, as well as how natural resources can be used in a sustainable manner to achieve India's sustainable development.

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