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Working Paper No. 1011

The Dynamics of Monthly Changes in US Swap Yields: A Keynesian Perspective

by

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ABSTRACT

John Maynard Keynes (1930) asserted that the central bank sways the long-term interest rate through the influence of its policy rate on the short-term interest rate. Recent empirical research shows that Keynes's conjecture holds for long-term Treasury yields in the United States. This paper investigates whether Keynes's conjecture also holds for the monthly changes in US longterm swap yields by econometrically modeling its dynamics using an autoregressive distributed lag (ARDL) approach. The econometric modeling reveals that there is statistically significant effect on the monthly changes in the Treasury bill rate on the monthly changes in swap yields of different maturity tenors after controlling for a host of macroeconomic and financial control variables. The findings from the econometric models that are estimated render a perspicacious Keynesian perspective on key policy questions and contemporary debates in macroeconomics and finance.

KEYWORDS: Interest Rate Swaps; Swap Yields; Short-Term Interest Rate; Monetary Policy; Federal Reserve; John Maynard Keynes

JEL CODES: E43; E50; E58; E60; G10; G12

INTRODUCTION

US-dollar-denominated interest rate swaps play a vital role in the US fixed income markets. As of second half of 2021, the notional value of outstanding US dollar interest rate swaps amounted to a bit more than \$134.6 trillion, while the gross market value of dollar-denominated interest swaps amounted to nearly \$1.8 trillion, according to the Bank for International Settlements (2022). While there is considerable discussion of many aspects of interest rate swaps, there is a clear lack of empirical literature on the modeling of dollar-denominated interest rate swap yields in terms of key macroeconomic and financial variables. This paper is a part of an effort, initiated in Akram and Mamun (2022), to address the lacuna regarding the macroeconomic and financial determinants of interest rate swap yields.

The existing literature in economics and quantitative finance have attempted to model swap yields as they relate to credit and liquidity conditions, rather than fundamental macroeconomic factors. The lack of any systematic attempts at empirically modeling the dynamics of swap yields based on macroeconomic and financial variables is a crucial gap in the existing literature. This paper specifically examines whether the short-term interest rate influences interest rate swap yields of different maturity tenors, after controlling for some key macroeconomic and financial factors. John Maynard Keynes's (1930) views on interest rate dynamics inspires an examination of the relationship between the swap yield and the short-term interest rate.

Keynes (1930) argued that the central bank's policy rate and other instruments of monetary policy exert an enormous influence over the long-term interest rate on gilt-edged securities via the short-term interest rate. His views on the relationship between the short-term interest rate and the long-term interest rate were based both on his theoretical insights on the workings of financial markets and investor behavior, as well as empirical regularities as documented in Riefler's (1930) statistical analysis of bond yields in the United States and Keynes's own observations of similar trends in the United Kingdom. In recent years, Akram and Li (2017, 2020), Deleidi and Levrero (2021), and Gabrisch (2022) have shown that Keynes's conjecture holds for long-term US Treasury yields.

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In this paper, the dynamics of dollar-denominated interest rate swap yields are examined to test whether the short-term interest rate has an influence on swaps yields of different maturity tenors after controlling for core inflation, the growth of industrial production, the change in the log of the equity price index, the log of the exchange rate, and the log of implied volatility of the equity market. By examining the relationship between the short-term interest rate and the swap yields of different maturity tenors, this paper will reveal whether Keynes's conjecture extends to not just to government bond yields but also to interbank interest rate swap yields.

Outline

This paper is organized as follows. Section II gives a brief primer on interest rate swaps. It also provides an overview of the literature on interest rate swaps. Section III develops a simple model that ties the interest rate swap yields to macroeconomic factors, including the short-term interest rate. Section IV describes the evolution of interest rate swap yields in the United States with reference to the macroeconomic milieu. Section V presents the data sources used in the paper, explains the variables, and displays summary statistics and unit root and stationary tests. Section VI reports the results of our econometric modeling of swap yields. Section VII concludes with reflections on the policy implications of the empirical results.

SECTION II: A PRIMER ON INTEREST RATE SWAPS AND A SHORT REVIEW OF THE LITERATURE

An interest rate swap is a simultaneous selling and purchasing of cash flows. An interest rate swap between two firms can be illustrated as follows: Firm A needs a \$1 million floating rate loan, whereas firm B needs a \$1 million fixed rate loan. However, firm B has a comparative advantage in a floating rate loan, whereas firm A has a comparative advantage in a fixed rate loan. Each firm borrows in the market where they have a comparative advantage. Hence, firm A borrows \$1 million in fixed rate, while firm B borrow \$1 million in floating rate. The firms then decide to exchange the interest payments with each other.

An interest rate swap constitutes an exchange of cash flow streams based on certain interest rates. Typically, for a plain vanilla swap, it is an exchange between a stream of fixed interest rate payments and a stream of floating interest rate payments, both in the same currency. The interest payments are based on the same notional principal. The floating rate is usually tied to some benchmark money market rate. The maturity of interest rate swaps varies but are usually between two to thirty years. The fixed rate payer is known as the buyer of the swaps, whereas the floating rate payer is known as the seller of the swap.

Interest rate swap terms are usually set such that the present value of the counterparty's payment is equal to the present value of the payment to be received. The initial value of the swap contract should be zero. The counterparty choosing to pay the fixed rate and the counterparty choosing to pay the floating rate each believe that they will gain. Their assumptions are based on their needs and their expectations of the level and changes in the interest rates during the tenor of the swap.

A Short Review of the Literature on Interest Rate Swaps

Bicksler and Chen (1986) and Wall and Pringle (1988) provide an economic analysis of interest rate swaps and their use in business and finance for liability transformation and asset transformation. Corb (2012) surveys the concepts underlying interest rate swaps and provides a lucid explanation of various topics concerning swaps, including risk characteristics, usage, and pricing. Swaps can be used for active liability management and hedging interest rate risks, as well as for speculating on directional views of future interest rates. Loeys (1985) has argued that interest rate swaps are useful for corporations and financial institutions for managing interest rate risks. Whittaker (1987) explains regulatory and policy issues concerning interest rate swaps. Kuprianov (1993) summarizes the importance of interest rate swaps in corporate finance. Studies of the use of interest rate swaps in corporations provide some valuable insights. For example, Chernenko and Faulkender (2011) examine nonfinancial firms' use of derivatives, such as interest rate swaps. They report that hedging of interest rate risk is concentrated among highinvestment firms, which is consistent with costly external finance. They also report that firms appear to use interest rate swaps to manage earnings and engage in speculation. Li and Mao (2003) show that fixed rate swap payers generally have lower credit ratings, a higher leverage ratio, a higher percentage of floating rate loans, and are more likely to use bank loans than

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floating rate swap payers. Duffie and Huang (1996) develop a model that ties the credit quality of a corporation to the swap yield, while Duffie and Singleton (1997) present a multifactor model of the term structure of interest rate swaps yields. Even though the literature on interest rate swaps is vast, the relationship between the short-term interest rate and the long-term swap yields has not been yet explored and econometrically modeled. Keynes's insight that the long-term interest rate moves largely in tandem with the short-term interest rates provides a solid foundation for scrutinizing the dynamics of long-term swap yields from a fundamental macroeconomic and financial perspective and econometrically modeling such dynamics.

SECTION III: A SIMPLE MODEL OF SWAP YIELD

A simple model of the swap yield is introduced here. The model draws on Akram (2021, 2022a, 2022b). Whereas those models operationalize Keynes's view that the short-term interest rate is the main driver of the long-term interest rate on government bonds, the model presented here relates the dynamics of the swap yield to the short-term interest rate and other key macroeconomic and financial variables.

The long-term interbank swap yield is S_{LT} . The short-term interest rate is r_{ST} . The central bank's policy rate is r_{CB} . The core inflation is π , while the central bank's core inflation target is $\bar{\pi}$. Z_1, Z_2, Z_3 are distinct Weiner processes. The parameters of the models are: $a_1, a_2, b_1, b_2, c_1, c_2$.

The model is expressed in the follow three equations:

$$dS_{LT}(t) = (a_1 r_{ST}(t) + a_2 \pi(t))dt + a_3 \sqrt{r_{ST}(t)} dZ_1$$
[1]

$$dr_{ST}(t) = b_1(r_{CB}(t) - r_{ST}(t))dt + b_2\sqrt{r_{ST}(t)}dZ_2$$
[2]

$$d\pi(t) = c_1(\overline{\pi} - \pi(t))dt + c_2\sqrt{\pi(t)}dZ_3$$
[3]

Equation [1] connects the dynamics of the interbank swap yield to the short-term interest rate, inflation, and a Weiner process adjusted by the short-term interest rate. Equation [2] tethers the dynamics of the short-term interest rate to: (i) the difference between the central bank's policy rate and (ii) the short-term interest rate and a Weiner process adjusted by the short-term interest rate. Equation [3] relates the dynamics of core inflation to: (i) the difference between the central bank's core inflation target and core inflation and (ii) a Weiner process adjusted by core inflation. Note that the each of the Weiner processes described in the above equations are independent and distinct from one another.

The above model can be readily extended to fuse other pertinent macroeconomic and financial variables, such as the growth of industrial production, the logarithm of the equity price index, the logarithm of the exchange rate, and the logarithm of the volatility of financial market, if these are regarded as having critical influence on the swap yield of different maturity tenors.

In the empirical part of the paper, the autoregressive distributed lag (ARDL) approach is used to econometrically model the dynamics of dollar-denominated swap yield and examine its relationship to the short-term interest rate, while controlling for the effects of other key macroeconomic and financial variables.

SECTION IV: THE EVOLUTION OF INTEREST RATE SWAP YIELDS AND ITS MACROECOMIC MILEU

Macroeconomic and financial market conditions have substantial influence on swap yields, as can be seen in this brief outline of US economic and financial variables. The Federal Reserve's monetary policy exerts considerable impact on swap yields via the short-term interest rate.

Figure 1 shows the evolution of outstanding US-dollar-denominated interest rate swaps. It gives both the notional value of outstanding interest rate swaps and the gross market value. The notional value of US-dollar-denominated swaps has risen from nearly \$14 trillion in 2000 to \$135 trillion in 2021, while the gross market value rose from \$0.4 trillion in 2000 to \$1.8 trillion

in 2021.¹ The notional value of outstanding dollar-denominated interest rate swaps rose steadily from 2000 to 2008, but it remained steady in the years following. However, the gross market value of dollar-denominated swaps peaked at \$10.4 trillion in 2008 and has steadily declined from 2011 to 2021.



Figure 1. The Evolution of Outstanding US-dollar-denominated Interest Rate Swaps, 2000–21

Figure 2 displays the evolution of swap yields of different maturity tenors, while figure 3 exhibits the coevolution 10-year swaps and the short-term interest rate. Swap yields declined from the beginning of 2011 to mid-2012 but rose from late 2012 to early 2014 followed by a decline until mid-2016. Swap yields began to decline and preceded the decline in the short-term interest rate from late 2016 to mid-2020. Swap yields from mid-2020 to the end of the period rose, while the short-term interest rate began to rise in late 2021.

¹ All figures are in current (nominal) US dollars.



Figure 2. The Evolution of Swap Yields in the United States, 2011M1–2022M3





Figure 4 traces the coevolution of 10Y swap yields and the core personal consumer expenditure (PCE) inflation. The figure would suggest that there is no tight connection between swap yields and core inflation.



Figure 4. The Coevolution of the 10-year Swap Yield and Core PCE Inflation, 2011M1–2022M3.

Figure 5 exhibits the evolution of the S&P500 index. The stock market has risen throughout the period under consideration. The rise in the S&P500 index was steady between 2011 and 2019. The S&P500 fell in March 2020 but it recovered in the following months and exceeded the previous peak by the late summer of 2020. The rise peaked at the end of 2021, and the S&P500 began to decline beginning in 2022.



Figure 5. The Evolution of the SP500 Index, 2011M1–2022M3

Figure 6 exhibits the evolution of the exchange rate of the US dollar against the euro. Here an increase in the exchange rate indicates a deprecation of the dollar, while a decrease in the exchange rate indicates an appreciation of the dollar.



Figure 6. The Evolution of the US-dollar Exchange Rate, EURUSD, \$/€, 2011M1–2022M3.

Figure 7 illustrates the evolution of the Chicago Board Options Exchange Volatility Index (VIX). The VIX is a measure of implied volatility. It is derived from real-time, mid-quote prices of call and put options of the S&P 500 index. The VIX rose notably in September of 2011 and remained high until the end of the year. The VIX stayed within a range of 10 to 25 from 2012 to 2019. In March 2020, with the onset of the global pandemic, the VIX rose sharply. It began to decline in the following month and normalized to less than 20 by March 2021. However, by the end of 2021, it began rising again and continued its increase as the Russian military operations started in February 2022.



Figure 7. The Evolution of the VIX, 2011M1–2022M3.

SECTION V: DATA DESCRIPTION, UNIT ROOT TESTS, AND STATIONARY TESTS

Table 1 provides a summary of the variables used in the paper. The first column lists the variable names. The second column gives the data description and the date range for the data. The third column enumerates the data frequency and indicates whether high frequency data have been converted to low frequency data. The final column provides the source of the data.

The short-term interest variables are 3- and 6-month Treasury bills. The interest rate swaps are of 2-, 5-, 10- and 30-year terms. Several controls variables, including inflation rates, industrial production index, two major exchange rates, two major US stock indices, and a market volatility measure, are also used. The financial control variables are in the (natural) $log_e(.)$ form.

Table 1. Summary	of the Data
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Variables	Data description, date range	Frequency	Sources
Short-term intere	est rates		
TB3M	Treasury bill, 3 month, %, bid yield at constant	Daily; converted to	Federal Reserve
	maturity, January 2011–March 2022	monthly	Board
TB6M	Treasury bill, 6 month, %, bid yield at constant	Daily; converted to	Federal Reserve
	maturity, January 2011–March 2022	monthly	Board
Long-term swap	rates		1
SWAP2Y	Interest rate swap, 2 year vs. 3 month float, %,	Daily; converted to	Tullet Prebon
	mid-rate, January 2011–March 2022	monthly	Information
SWAP5Y	Interest rate swap, 5 year, vs. 3 month float,	Daily; converted to	Tullet Prebon
	%, mid-rate, January 2011–March 2022	monthly	Information
SWAPIOY	Interest rate swap, 10 year vs. 3 month float,	Daily; converted to	Tullet Prebon
	%, mid-rate, January 2011 — March 2022	monthly	Information
SWAP30Y	Interest rate swap, 30 year, vs. 3 month float,	Daily;	Tullet Prebon
	%, mid-rate, January 2011 — March 2022	converted to	Information
In fl at an		montniy	
Injunion	Demonstration and ditem loss for all 9	Manthla	Dunnen of Door on in
COREFCE	energy: chain price index, seasonally adjusted	Monuny	Analysis
	% change $y/y = 2012 - 100$ January 2011		Allarysis
	March 2022		
CORECPI	Consumer price index all items less food and	Monthly	Bureau of Labor
conderr	energy, not seasonally adjusted. % change.	Wollding	Statistics
	v/v, 1982-1984 = 100. January 2011–March		
	2022		
Economic activit	y v		
IPYOY	Industrial production, index, % change, y/y,	Monthly	Federal Reserve
	seasonally adjusted, 2017 = 100, January		Board
	2011–March 2022		
Financial variab	les		
YEN	Exchange rate, Japanese yen per US dollar,	Daily; converted to	Federal Reserve
	¥/\$, January 2011–March 2022	monthly	Board
EURO	Exchange rate, US dollar per euro, \$/€,	Daily; converted to	Federal Reserve
	January 2011–March 2022	monthly	Board
DIJA	Stock price index, Dow Jones: 30 industrial	Daily; converted to	Wall Street Journal
	stocks, average price close, January 2011–	monthly	
	March 2022		
SP500	Standard & Poor's 500 composite index,	Daily; converted to	Standard and Poor's
VIV	January 2011–March 2022	monthly Deilen commune 14	
VIX	Unicago Board of Exchange (UBUE), market	Daily; converted to	wall Street Journal
	volatility index. January 2011–March 2022	monthly	

The summary statistics of all the variables are reported in table 2A and table 2B. Table 2A displays the summary statistics of the variables, while table 2B provides the summary statistics of the first differences of the variables. The longer-term swap rates and stock market indices are not normally distributed. However, all variables are normally distributed when converted to first-difference.

	Obs	Mean	Std. Dev.	Max	Min	Skewness	Kurtosis	JB	Probability
SWAP2Y	134	1.05	0.77	3.05	0.20	1.13	3.23	28.56	0.00
SWAP5Y	134	1.56	0.66	3.14	0.31	0.31	2.79	2.41	0.30
SWAP10Y	134	2.13	0.67	3.68	0.60	-0.18	2.74	1.09	0.58
SWAP30Y	134	2.58	0.77	4.42	0.86	-0.04	2.93	0.07	0.97
TB3M	134	0.55	0.78	2.45	0.01	1.35	3.28	41.38	0.00
TB6M	134	0.61	0.80	2.54	0.04	1.30	3.17	37.67	0.00
COREPCE	134	1.83	0.76	5.40	0.91	2.91	11.84	625.56	0.00
CORECPI	134	2.13	0.85	6.41	0.95	2.95	12.66	715.30	0.00
IPYOY	134	0.96	4.31	17.81	-17.69	-0.57	8.76	192.78	0.00
LNYEN	134	4.64	0.13	4.82	4.34	-1.08	2.97	25.82	0.00
LNEURO	134	0.19	0.09	0.37	0.05	0.47	1.96	10.85	0.00
LNDIJA	134	9.89	0.33	10.49	9.32	0.11	1.94	6.54	0.04
LNSP500	134	7.71	0.36	8.45	7.07	0.14	2.28	3.36	0.19
LNVIX	134	2.83	0.31	4.06	2.32	1.01	4.54	35.92	0.00

Table 2.A: Summary Statistics of the Variables

Note: $LN = Natural \log = Log_e(.)$.

Table 2.B: Summary Statistics of the First Differences of the Variables

	Obs	Mean	Max	Min	Std. Dev.	Skewness	Kurtosis	JB	Probability
ΔSWAP2Y	133	0.01	0.47	-0.80	0.13	-1.66	13.91	720.68	0.00
∆SWAP5Y	133	0.00	0.39	-0.66	0.16	-0.67	4.95	30.95	0.00
∆SWAP10Y	133	-0.01	0.42	-0.66	0.17	-0.55	4.73	23.18	0.00
∆SWAP30Y	133	-0.02	0.37	-0.69	0.17	-0.88	5.23	44.68	0.00
ΔTB3M	133	0.00	0.36	-1.24	0.13	-6.24	59.76	18714.27	0.00
∆TB6M	133	0.00	0.31	-1.21	0.13	-6.36	62.96	20820.06	0.00
ΔCOREPCE	133	0.03	1.11	-0.79	0.17	1.45	17.06	1142.47	0.00
∆CORECPI	133	0.04	1.31	-0.66	0.21	2.32	14.74	883.01	0.00
ΔΙΡΥΟΥ	133	0.02	15.96	-12.35	2.23	1.60	28.90	3774.27	0.00
ΔLNYEN	133	0.00	0.07	-0.04	0.02	0.79	4.41	24.90	0.00
ΔLNEURO	133	0.00	0.04	-0.06	0.02	-0.34	3.65	4.98	0.08
ΔLNDIJA	133	0.01	0.11	-0.23	0.04	-2.44	17.95	1370.40	0.00
ΔLNSP500	133	0.01	0.09	-0.21	0.03	-2.47	16.58	1157.36	0.00
ΔLNVIX	133	0.00	1.08	-0.37	0.20	1.69	9.77	317.60	0.00

Note: Δ is the first-difference of the variable

Table 3A displays the results of the unit root tests, stationary, and the stationary tests for these variables. The unit root tests are conducted using the automated Dickey-Fuller (ADF) tests (Dickey and Fuller 1979, 1981) and the Phillips Perron tests (Phillips and Perron 1988), while the stationarity tests are conducted using Kwiatkowski, Phillips, Schmidt, and Shin (KPSS) tests (Kwiatkowski et al. 1992). The test results are mixed in table 3A, with most variables nonstationary at the level.

Table J.A.		anu Statiu	marity 103	sis of the v	arrapics				
	ADF Unit Ro	oot Tests (H ₀ : No	nstationary)	PP Unit Ro	ot Tests (H ₀ : Nor	nstationary)	KPSS Tests (H ₀ : Stationarity)		
	None	Intercept	Trend	None	Intercept	Trend	Intercept	Trend	
SWAP2Y	-0.28	- 1.54	- 1.91	-0.48	- 1.47	- 1.72	0.35*	0.17**	
SWAP5Y	- 0.91	-2.14	-2.08	-0.77	-2.05	- 1.94	0.14	0.14*	
SWAP10Y	- 1.35	- 2.92**	-2.66	-1.18	- 2.49	-2.34	0.45*	0.08	
SWAP30Y	- 1.53	-2.82*	- 3.03	- 1.45	-2.36	- 2.53	0.89***	0.05	
TB3M	-0.88	- 1.22	- 1.12	- 1.05	- 1.38	- 1.33	0.45*	0.16**	
TB6M	-0.77	- 1.28	- 1.36	-0.96	- 1.45	- 1.53	0.46*	0.17**	
COREPCE	1.05	0.01	-0.48	2.27	2.18	1.40	0.46*	0.15**	
CORECPI	1.13	-0.72	- 1.50	2.05	1.62	0.86	0.51**	0.14*	
IPYOY	-2.38**	-2.41	-2.40	- 3.01***	- 3.10**	- 3.04	0.16	0.07	
LNYEN	1.13	- 1.54	- 1.69	1.14	- 1.51	- 1.62	0.80***	0.28***	
LNEURO	- 1.53	-1.74	- 2.13	- 1.34	- 1.65	-2.17	0.85***	0.20**	
LNDIJA	2.67	- 0.36	- 3.45**	4.15	-0.05	- 3.43**	1.42***	0.06	
LNSP500	3.47	- 0.05	-2.78	3.96	0.16	- 2.68	1.39***	0.11	
LNVIX	- 0.19	- 3.73***	- 3.80**	0.30	- 3.57***	- 3.64**	0.28	0.23***	

 Table 3.A: Unit Root and Stationarity Tests of the Variables

Note: Significance levels: *** for 1 percent, ** for 5 percent and * for 10 percent

Table 3B presents the unit root and the stationarity tests for the first differences of the variables. All the ADF unit root tests indicate that the null hypothesis of nonstationarity can be rejected at the 1 percent level of significance for the first differences of all variables. The PP tests also reject the null hypothesis of nonstationary at the 1 percent level for all first-differenced variables. The KPSS tests show that the null hypothesis of stationarity cannot be rejected for the first differences of these variables (except for the two measurements for inflation).

	ADF Unit R	oot Tests (H ₀ : No	onstationary)	PP Unit Ro	ot Tests (H ₀ : No	nstationary)	KPSS Tests (H ₀ : Stationarity)		
	None	Intercept	Trend	None	Intercept	Trend	Intercept	Trend	
∆SWAP2Y	- 6.29***	- 6.29***	- 6.32***	- 6.49***	- 6.49***	-6.68***	0.10	0.10	
∆SWAP5Y	- 8.32***	- 8.29***	-8.38***	-8.47***	-8.44***	-8.52^{***}	0.12	0.10	
∆SWAP10Y	- 8.30***	- 8.29***	- 8.37***	- 8.32***	- 8.29***	- 8.37***	0.12	0.08	
∆SWAP30Y	-7.92***	- 7.95***	-8.01^{***}	- 7.92***	- 7.95***	-8.01^{***}	0.11	0.06	
∆TB3M	- 9.69***	- 9.66***	- 9.64***	- 9.96***	- 9.93***	- 9.91***	0.14	0.12*	
∆TB6M	- 8.43***	- 8.41***	- 8.37***	-8.71***	- 8.69***	-8.65^{***}	0.11	0.10	
<i>ACOREPCE</i>	-4.68***	-4.80***	-5.08***	-7.15^{***}	- 7.31***	- 7.43***	0.45*	0.21**	
<i>ACORECPI</i>	- 3.13***	- 3.30**	- 3.41*	- 5.51***	- 5.49***	- 5.29***	0.48**	0.22***	
ΔΙΡΥΟΥ	- 5.57***	- 5.53***	- 5.57***	- 9.18***	-9.12***	- 9.20***	0.10	0.05	
ΔLNYEN	- 8.29***	- 8.38***	- 8.37***	- 8.33***	- 8.34***	- 8.32***	0.12	0.08	
ALNEURO	- 8.74***	- 8.77***	-8.74***	- 8.90***	- 8.90***	-8.86^{***}	0.06	0.05	
	_	-	-	-	-	-	0.07	0.05	
ALNDIJA	11.18***	11.71***	11.68***	11.18***	12.59***	12.62***			
	_	- 9.70***	-9.70***	_	_	_	0.08	0.05	
Δ LNSP500	10.34***			10.31***	11.16***	11.24***			
						$-\overline{18.46^{***}}$	0.18	0.11	
Δ LNVIX	12.07***	12.03***	12.00***	16.91***	16.87***				

Table 3.B: Unit Root and Stationarity Tests of the First Differences of the Variables

Note: Significance levels: *** for 1 percent, ** for 5 percent and * for 10 percent

SECTION VI: ECONOMETRIC MODELS AND EMPIRICAL FINDINGS

The autoregressive distributed lag approach (ARDL) is applied for the estimation of short-term relationships between the short-term interest rate and the swap rate. The basic form of an ARDL regression model is:

$$\Delta y_t = \beta_0 + \beta_1 \Delta y_{t-1} + \dots + \beta_k \Delta y_{t-p} + \alpha_0 \Delta x_t + \alpha_1 \Delta x_{t-1} + \dots + \alpha_q \Delta x_{t-q} + \varepsilon_t$$

where y and x are the swap yield and Treasury bill rate respectively and ε_t is a random "disturbance" term. The main results with different swap term rates and 3-months Treasury bills rate are presented in table 4. The core PCE inflation and growth in industrial production at their first difference are presented as controls in the first models for all swap rates. Then the model is expanded to other the control variables by adding a measurement of stock return (Δ LNSP500), the percent change in the euro–dollar exchange rate (Δ LNEURO), and the percent change in market volatility (Δ LNVIX). The control variables aim to adjust for any macroeconomic impact on the interbank swap rates.

The ARDL approach is used to econometrically model the dynamics of swap yields rather than another approach for several important reasons. First, the autoregressive conditional heteroscedasticity Lagrange multiplier (ARCH-LM) tests on the ordinary least squares (OLS) model of swap and Treasury rates are conducted and the results of these tests indicated that there is no ARCH effect present in US Treasury bills and swap rates except for 2-year swap rates (see Appendix A for the detailed results). Second, the traditional vector autoregressive (VAR) and vector error correction models (VECM) are also applied. However, the Granger causality tests did not show any two-way directional relationship between swap and Treasury rates. Similarly, the VECM did not reveal any converging models for the short-term interest rates and swap yields of different maturity tenors. Lastly, tests for structural breaks did not yield any significant results. Based on the evaluation of these findings, the ARDL approach appears to be most germane for the analysis of the dynamics of swap yields.

The 3-month Treasury rate at the first difference has a positive relationship with all the different terms of the swap rates. The size of the impact is diminished the higher the maturity tenor of the swaps. This finding supports the hypothesis that Keynesian inferences on the influence of the short-term interest rate can be extended to interbank swap rates. Furthermore, the short-term Treasury rate impacts the swap rates with a one-period lag. However, the impact is negative and pronounced for higher term lengths. Keynes did not remark on the lead-lag structure of the influence of the short-term interest rate on the long-term interest rate. However, these findings do not in any manner contradict the Keynesian view that the short-term interest rate has an important and decisive influence on swap rates. Furthermore, the lagged swap rates are very good predictors of current rates.

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	∆SWAP2Y	∆SWAP2Y	∆SWAP5Y	∆SWAP5Y	∆SWAP10Y	∆SWAP10Y	∆SWAP30Y	∆SWAP30Y
			Main	Equation				
ΔΤΒ3Μ	0.59***	0.53***	0.57***	0.43***	0.54***	0.37***	0.55***	0.39***
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
Δ ΤΒ3Μ(-1)	-0.17**	-0.16	-0.17	<u>``</u>	-0.21*		-0.27***	-0.16**
	(0.04)	(0.12)	(0.10)		(0.06)		(0.00)	(0.05)
Δ SWAP_Y(-1)	0.34**	0.39***	0.24**	0.25***	0.29***	0.29***	0.36***	0.37***
	(0.01)	(0.00)	(0.01)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
△COREPCE	0.07	0.05	0.06	0.04	-0.009	-0.04	-0.06	-0.08
	(0.33)	(0.49)	(0.52)	(0.72)	(0.92)	(0.69)	(0.50)	(0.38)
ΔΙΡΥΟΥ	-0.002	-0.001	0.0001	-0.003	0.002	-0.001	0.002	-0.0003
	(0.56)	(0.69)	(0.99)	(0.63)	(0.66)	(0.81)	(0.61)	(0.96)
∆LNSP500		-0.15		0.55		0.62		0.50
		(0.83)		(0.54)		(0.50)		(0.55)
ΔLNEURO		-0.08		0.09		0.69*		1.36*
		(0.87)		(0.90)		(0.37)		(0.07)
ΔLNVIX		-0.10		-0.09		-0.14		-0.17
		(0.30)		(0.51)		(0.37)		(0.26)
Intercept	0.001	0.003	-0.01	-0.01	-0.01	-0.01	-0.02	-0.01
	(0.89)	(0.77)	(0.62)	(0.54)	(0.49)	(0.45)	(0.42)	(0.43)
			Model I	nformation	L			
Obs	132	132	132	132	132	132	132	132
Adj R ²	0.51	0.51	0.29	0.30	0.25	0.27	0.29	0.35
AIC	- 1.89	- 1.87	- 1.11	- 1.12	-0.90	- 0.93	-0.98	-1.05
			Diagno	ostic Tests				
Joint significance	27.94	17.91	11.61	8.88	9.53	8.06	11.92	10.03
F-Test	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
Serial correlation	1.98	1.98	1.93	1.91	1.95	1.92	1.96	2.00
Durbin-Watson								
Statistic	0.65	0.67	0.40	0.01	0.01	0.00	0.00	0.00
Serial correlation	0.65	0.67	0.42	0.21	0.24	0.22	0.09	0.20
Breusch-Godfrey	(0.52)	(0.52)	(0.66)	(0.81)	(0.79)	(0.80)	(0.90)	(0.82)
LNI Test	1.02	1.12	0.40	1.50	0.70	2.51	0.97	2.24
Heteroskedasticity	1.02	1.13	(0.85)	1.50	(0.79)	2.51	(0.87)	2.24
Dreusch-Pagan-	(0.41)	(0.55)	(0.83)	(0.17)	(0.36)	(0.02)	(0.30)	(0.05)
Normality test	2.62	1 / 9	0.58	0.15	0.41	0.87	1.01	1 1 9
Normanty test	(0.16)	1.40	(0.38)	(0.03)	(0.41)	(0.67)	(0.60)	(0.55)
statistic	(0.10)	(0.40)	(0.73)	(0.93)	(0.01)	(0.05)	(0.00)	(0.55)
Stability diagnostic	1 59	2.07	1 49	2 53	1.62	3.81	0.61	3.45
Ramsev RESET	(0.21)	(0.13)	(0.23)	(0.08)	(0.20)	(0.02)	(0.54)	(0.03)
Test	(0.21)	(0.15)	(0.23)	(0.00)	(0.20)	(0.02)		(0.05)

Table 4: ARDL (p, q) Model (with \triangle TB3M)

Note: All vars are in diff, *p*-values are in parenthesis. BG LM and Ramsey RESET tests are with 2 lags.

Most of the controls variable, such as core PCE inflation, the growth of the industrial production, the percentage change in the stock price index, the euro–dollar exchange rate, and the percentage change in market volatility, did not yield any meaningful relationship with the swap rate, but it is still useful to examine whether these have any impact whatsoever.

Other diagnostic tests are also presented in table 4. There are no serial correlations in these models based on (1) the Durbin Watson test and (2) Breusch and Godfrey LM test. The Jarque-Bera tests indicate the error terms have the skewness and kurtosis matching a normal distribution. The regression errors are homoscedastic (except for two cases). In addition, the Ramsey RESET tests indicate the general specifications of the models are correct. The cumulative sum (CUMSUM) and the cumulative sum of squares (CUMSUMQ) tests for parameter stability are presented in appendix B. These tests show that the time-series models developed are stable over time and without any structural breaks. The Ljung-Box Q-statistics and their p-values as part of the correlogram for these models are provided in appendix C. The Q-statistics consistently failed to reject the null hypothesis of no autocorrelation.

A robustness check with the impact of the 6-month Treasury bill rate (instead of the 3-month Treasury bill rate) on the different term swap rates is undertaken in appendix D. A set of different control variables, such as the first difference of core CPI for an inflation measurement, stock return of the Dow-Jones industrial average index, and percent change in the yen–dollar exchange rate, are also utilized to check whether the results hold. The regression using the 6-month Treasury bill rate showed a similar positive relationship with the swap rates as shown in the table 4 results. Lastly, the ARCH-LM tests of the models in table 4 are displayed in appendix E. Similar to the ARCH-LM tests in the OLS models (appendix A), there is no evidence of heteroscedastic residuals (i.e., does not exhibit ARCH effects) except for 2-year swap rates.

SECTION VII: CONCLUSION

The empirical findings of the paper have bearing on macroeconomic theory and policy, particularly concerning monetary policy and financial market regulations, asset allocation, and risk management. The findings clearly show that the Federal Reserve's decisions regarding the fed funds target rate and other monetary policy matters exert a marked effect on the swap yields of different maturity tenors through the monthly changes in the short-term interest rate. An increase (decrease) in the short-term interest rate tends to be associated with a concomitant increase (decrease) in the long-term swap yield. The effect is most pronounced at the front end of the swap yield curve. Monthly changes in most control variables have hardly any discernable effect on the monthly changes in the swap yield. This supports that case that the short-term interest rate is the key driver of long-term swap yields of different maturity tenors after controlling for assorted factors. These empirical findings bolster the case that the Federal Reserve does and can exert substantive influence on financial markets and in particular its actions sway the pricing of fixed income securities and derivatives, including interest rate swaps through monetary policy actions (Bindseil 2004; Fullwiler [2008] 2017).

The results obtained from the empirical modeling of swap yields shows the important influence of the short-term interest rate on the swap yields of different maturity tenors, after controlling for several key macroeconomic and financial variables, such as core inflation, the growth of industrial production, the log of the equity price index, the log of the exchange rate, and the log of the implied volatility of the equity index. The results suggest that the Federal Reserve's monetary policy, through the effects of the feds funds target rate (range) on the short-term interest rate and other actions, influence the dollar-denominated swap yields. This is a substantive finding given the wide-ranging influence of swap yields on financial intermediation, the banking industry, nonfinancial corporate borrowing and lending, capital markets, financial institutions, financial stability, and the real economy.

The empirical findings extends John Maynard Keynes's conjecture that the central bank influences not just the long-term interest rate on government securities but also the long-term swap yield. It reveals the ability of the central bank to influence the interbank interest rate, borrowing and lending rates, financial intermediation, and the financial system. The findings concerning the dynamics of US-dollar-denominated swap yields is not just relevant for US financial markets but also for understanding the dynamics of swap yields in other major financial markets. These findings could advance the empirical modeling of swap yields based on fundamental macroeconomic and financial factors.

This paper fulfils a critical lacuna in the empirical literature on the swap yield, as most often in the existing literature in economics and quantitative finance, the swap yield is empirically modeled merely as a function of the Treasury yield of the same maturity tenor and a spread,

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rather than the underlying macro and financial factors. The approach used here can prove to be fruitful in further investigations of swap yield dynamics.

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Models	∆SWAP2Y	∆SWAP5Y	∆SWAP10Y	∆SWAP30Y	∆SWAP2Y	∆SWAP5Y	∆SWAP10Y	∆SWAP30Y
		ΔΤΙ	33M			ΔΤΙ	B6M	
Lags				Pane	l One			
1	3.79	2.95	5.64	4.69	10.61	0.11	1.74	4.21
	(0.05)	(0.09)	(0.02)	(0.03)	(0.00)	(0.74)	(0.19)	(0.04)
4	2.30	1.05	2.01	1.66	3.39	0.14	0.76	1.88
	(0.06)	(0.39)	(0.09)	(0.16)	(0.01)	(0.97)	(0.55)	(0.12)
8	2.41	0.89	1.07	0.49	2.81	0.48	0.83	0.48
	(0.02)	(0.53)	(0.39)	(0.86)	(0.01)	(0.87)	(0.58)	(0.86)
12	1.89	0.97	0.93	0.48	2.28	0.73	0.73	0.40
	(0.04)	(0.48)	(0.52)	(0.92)	(0.01)	(0.72)	(0.72)	(0.96)
				Pane	Two			
1	3.83	0.73	2.52	2.27	1.62	0.00	1.19	2.38
	(0.05)	(0.39)	(0.11)	(0.13)	(0.20)	(0.94)	(0.27)	(0.12)
4	2.28	0.40	1.13	1.24	1.51	0.02	0.54	1.07
	(0.06)	(0.81)	(0.36)	(0.30)	(0.20)	(0.99)	(0.70)	(0.37)
8	2.45	0.55	1.01	0.95	1.46	0.44	0.87	0.81
	(0.02)	(0.82)	(0.43)	(0.47)	(0.18)	(0.89)	(0.55)	(0.60)
12	1.93	0.69	0.93	0.97	1.66	0.65	0.88	0.96
	(0.04)	(0.75)	(0.52)	(0.48)	(0.09)	(0.80)	(0.57)	(0.49)

Appendix A: ARCH LM test

Note: OLS model includes the change in the short-term interest rate (Δ TB3M, Δ TB6M) and the controls (namely Δ COREPCE and Δ IPYOY in panel one and Δ COREPCE, Δ IPYOY, Δ LNSP500, Δ LNEURO, and Δ LNVIX in panel two). *p*-values are in parenthesis.



Appendix B.1: CUMSUM for ARDL (p, q) Model (with Δ **TB3M)**

Note: ARDL (p, q) models include the change in the short-term interest rate (Δ TB3M) and the controls (namely Δ COREPCE and Δ IPYOY in the base model and Δ COREPCE, Δ IPYOY, Δ LNSP500, Δ LNEURO, and Δ LNVIX in the extended model).



Appendix B.2: CUMSUMSQ for ARDL (p, q) model (with Δ TB3M)



Note: ARDL (p, q) models include the change in the short-term interest rate (Δ TB3M) and the controls (namely Δ COREPCE and Δ IPYOY in the base model and Δ COREPCE, Δ IPYOY, Δ LNSP500, Δ LNEURO, and Δ LNVIX in the extended model).

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob*
. 0		1 -0.0	37 -0.037	0.1875	0.665
· 🗓 ·	j j	2 0.0	80 0.079	1.0613	0.588
ı İ L ı	j j	3 0.0	55 0.061	1.4774	0.687
i ≬ i	111	4 0.0	09 0.007	1.4891	0.829
()	111	5 -0.0	11 -0.020	1.5062	0.912
ı 🗐 i	ıþ	6 0.0	93 0.088	2.7127	0.844
i 🔲 i	י 🗖 י	7 -0.0	91 -0.085	3.8972	0.792
i ≬ i	I I	8 0.0	04 -0.016	3.8990	0.866
i 🏚 i	ן ו 🏚 י	9 0.0	52 0.057	4.2834	0.892
· □ ·	ı D ı	10 -0.1	23 -0.114	6.4905	0.773
	I I	11 -0.0	12 -0.025	6.5111	0.837
i 🏚 i	ן ום י	12 0.0	63 0.069	7.0874	0.852
1 1 1	ן ום ו	13 0.0	20 0.055	7.1493	0.894
1 1 1	1 1	14 0.0	23 0.012	7.2281	0.926
i 🏚 i	ן ום	15 0.0	66 0.046	7.8790	0.929
1 ()	1 1	16 -0.0	09 0.017	7.8908	0.952
ים,	I I ⊡ I	17 -0.0	78 -0.114	8.8281	0.946
i 🏚 i	1 1	18 0.0	32 0.004	8.9914	0.960
יםי	111	19 -0.0	49 -0.016	9.3741	0.967
י 🛄 י	וםי ו	20 -0.0	74 -0.087	10.237	0.964
i 🏚 i	ן ום	21 0.0	63 0.052	10.867	0.965
1 1 1	1 1	22 -0.0	43 0.001	11.160	0.972
i ∎ i	1 1	23 -0.0	16 0.005	11.199	0.981
יםי	ים י	24 -0.0	65 -0.093	11.883	0.981
ים י	יםי	25 -0.0	70 -0.055	12.705	0.980
i 🏚 i	י 🗗 י	26 0.0	42 0.074	13.001	0.984
i 🖡 i	111	27 0.0	25 -0.015	13.108	0.989
1 1 1	111	28 -0.0	24 -0.010	13.206	0.992
יםי	יםי	29 -0.0	58 -0.061	13.793	0.992
· 🗖 ·		30 -0.1	16 -0.146	16.107	0.982
• U	1 1 1 1	31 -0.0	27 -0.007	16.238	0.986
י 🗖 י	י ד וי	32 -0.1	37 -0.122	19.548	0.959
ים י	יםי	33 -0.0	73 -0.060	20.505	0.956
1 ()	1 1	34 -0.0	17 -0.010	20.560	0.966
1 D 1	· ·	35 0.0	61 0.089	21.233	0.968
· ت ا ا	ı [ı	36 -0.0	88 -0.057	22.657	0.959

Appendix C: Correlogram – Q -Stat ARDL (p, q) models with Δ TB3M Table C.1: Δ SWAP2Y = Φ^1 (C, Δ TB3M, Δ COREPCE, Δ IPYOY)

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
		1	0.022	0.022	0.0642	0.800
1 j 1	j i j i	2	-0.018	-0.018	0.1062	0.948
ı 👝 i	j j j j j j j j j j j j j j j j j j j	3	0.142	0.143	2.8700	0.412
i 🖬 i	j i g i	4	-0.078	-0.087	3.7171	0.446
i 🏚 i		5	0.057	0.069	4.1664	0.526
r 🏚 i		6	0.114	0.088	5.9912	0.424
ا 🛛 ا	ו 🛛 י	7	-0.084	-0.068	6.9965	0.429
1 1 1	I I	8	0.024	0.011	7.0768	0.528
1 () 1	וםי	9	-0.032	-0.057	7.2216	0.614
, D		10	-0.069	-0.033	7.9101	0.638
i 🎙 i	1 1	11	0.039	0.013	8.1276	0.702
i 🏚 i	ן וים	12	0.055	0.067	8.5802	0.738
1 1 1	1 1	13	-0.020	-0.004	8.6395	0.800
i 🏚 i	ן וףי	14	0.062	0.050	9.2187	0.817
· 🗐 ·	ı -	15	0.112	0.120	11.127	0.744
1 0 1	וויי	16	-0.049	-0.048	11.499	0.778
1 0 1	ים,	17	-0.058	-0.089	12.021	0.799
· 🛛 ·		18	0.059	0.035	12.553	0.817
<u> </u>	□ '	19	-0.186	-0.181	17.943	0.526
101	וויי	20	-0.028	-0.032	18.068	0.583
1 D 1	1 1	21	0.046	0.023	18.402	0.623
ا ا	ן ינןי	22	-0.109	-0.030	20.327	0.563
· • •		23	-0.015	-0.023	20.362	0.620
	'Щ'	24	-0.113	-0.114	22.467	0.551
		25	-0.166	-0.095	27.037	0.354
1	""] '	26	-0.056	-0.124	27.562	0.380
		27	-0.043	-0.031	27.870	0.418
		28	-0.025	-0.001	27.980	0.466
		29	-0.006	-0.035	27.986	0.519
		30	-0.053	-0.017	28.469	0.546
		31	-0.179	-0.154	34.090	0.321
		32	-0.048	-0.050	34.500	0.349
· u ·		133	-0.035	-0.062	34./1/	0.386
· · · ·		134	-0.020	0.018	34.039	0.428
		135	0.161	0.167	39.584	0.273
· ų ·	ו ישי	130	-0.046	-0.061	39.981	0.298

Table C.2: Δ SWAP5Y = $\Phi^1(C, \Delta$ TB3M, Δ COREPCE, Δ IPYOY)

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
111	 .	1	0.018	0.018	0.0452	0.832
i 🖞 i	ן ון ו	2	-0.031	-0.031	0.1758	0.916
i 🏚 i		3	0.114	0.115	1.9504	0.583
ا 🛛 ا	י 🗖 י	4	-0.092	-0.099	3.1243	0.537
1 D 1		5	0.081	0.096	4.0292	0.545
· 🗖 ·		6	0.104	0.080	5.5370	0.477
ا 🛛 ا	I I	7	-0.110	-0.091	7.2422	0.404
1 1 1	111	8	0.003	-0.012	7.2435	0.511
1 (1	I I	9	-0.022	-0.035	7.3156	0.604
1 (1	1 1	10	-0.016	0.018	7.3537	0.692
1 D 1	ן וףי	11	0.064	0.031	7.9604	0.717
1 1	1 1	12	-0.007	0.003	7.9669	0.788
ا ا ا	ן ו ב ו ו	13	-0.096	-0.083	9.3439	0.747
i 🏚 i	ן וףי	14	0.059	0.054	9.8594	0.772
i 🏚 i	ן וים	15	0.082	0.090	10.872	0.762
. .	I I	16	-0.043	-0.046	11.156	0.800
. .	וםי	17	-0.032	-0.063	11.316	0.840
т р т	ן וףי	18	0.035	0.055	11.510	0.871
— 1		19	-0.178	-0.164	16.469	0.626
. .	I I I I	20	-0.025	-0.046	16.568	0.681
I.↓ I	1 1	21	0.022	0.002	16.642	0.733
I 🗐 I	I I I I	22	-0.101	-0.045	18.270	0.690
. .	יםי	23	-0.039	-0.062	18.523	0.729
ا ا ا	וםי	24	-0.083	-0.071	19.656	0.716
I 🔲 I	וםי	25	-0.111	-0.065	21.681	0.654
י 📕 י	l I	26	-0.083	-0.146	22.838	0.642
1 0 1	I I I I	27	-0.058	-0.031	23.396	0.664
1 ()	1 1	28	-0.018	0.001	23.451	0.710
i 🖡 i	1 1	29	0.034	0.017	23.648	0.746
1 Q (ן יעי	30	-0.046	-0.040	24.013	0.771
ا ا	l ⊡ '	31	-0.129	-0.118	26.931	0.676
1 I 1	1 1	32	0.015	-0.018	26.970	0.719
1 1 1	י ם י	33	-0.064	-0.084	27.697	0.728
1 I 1	ן וויי	34	-0.007	0.027	27.705	0.769
· 🗩	ı 	35	0.165	0.139	32.655	0.582
· 🛛 ·	I I	36	-0.050	-0.044	33.108	0.607

Table C.3: Δ SWAP10Y = $\Phi^1(C, \Delta$ TB3M, Δ COREPCE, Δ IPYOY)

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
111	 .	1	0.012	0.012	0.0196	0.889
ı di i	j ingi	2	-0.036	-0.037	0.2007	0.905
ı (İ) ı	j j	3	0.071	0.072	0.8872	0.829
ı 🔲 i	i n d in	4	-0.125	-0.129	3.0439	0.551
ı 🗍 ı	j j	5	0.067	0.080	3.6780	0.597
ı (İD)		6	0.089	0.072	4.7995	0.570
1 0 1	ון ו	7	-0.063	-0.045	5.3612	0.616
1 i 1	1 1 1	8	0.005	-0.012	5.3647	0.718
1 1 1	I I	9	0.002	0.005	5.3652	0.801
1 1 1	I I	10	-0.007	0.015	5.3731	0.865
r 🏚 r		11	0.078	0.056	6.2736	0.855
1 0 1	I I	12	-0.032	-0.038	6.4271	0.893
ا ا	l ⊨ ∎ i	13	-0.121	-0.110	8.6125	0.802
i 🏚 i	ן וףי	14	0.046	0.043	8.9320	0.835
i 🖡 i	ן ומי	15	0.025	0.037	9.0287	0.876
. ()	ון ו	16	-0.040	-0.045	9.2755	0.902
i 🖡 i	יםי	17	-0.015	-0.054	9.3113	0.930
1 1	ן ום ד	18	0.006	0.042	9.3162	0.952
· 🔲 ۱	ים י	19	-0.117	-0.102	11.456	0.908
1 0 1	יםי	20	-0.056	-0.080	11.944	0.918
1 Q 1	1 1	21	0.028	0.020	12.073	0.938
· 🖪 ·	וויי	22	-0.067	-0.046	12.786	0.939
· 🛛 ·	ים י	23	-0.077	-0.094	13.741	0.934
· U ·	יםי	24	-0.058	-0.059	14.298	0.940
1	I <u>I</u> I	25	-0.086	-0.068	15.513	0.929
۱ ۵ - ۱	™	26	-0.094	-0.129	16.993	0.909
1	1 1	27	-0.020	-0.024	17.062	0.930
		28	0.019	0.033	17.125	0.946
· 🛛 ·		29	0.058	0.045	17.704	0.950
1	ן ים י	30	-0.065	-0.080	18.446	0.951
יוןי	ן יעי	31	-0.057	-0.035	19.013	0.955
· 🛛 ·		32	0.058	0.035	19.605	0.958
' U ''	'■' '	33	-0.102	-0.112	21.462	0.939
' <u> </u> '	' ['	34	0.016	0.020	21.509	0.953
' P '	' □ '	35	0.137	0.118	24.950	0.896
<u> </u>	וויי	36	-0.070	-0.065	25.858	0.894

Table C.4: Δ SWAP30Y = $\Phi^1(C, \Delta$ TB3M, Δ COREPCE, Δ IPYOY)

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
		1	-0.040	-0.040	0.2150	0.643
. j	i ingi	2	0.082	0.080	1.1260	0.570
ı (D)	i j	3	0.079	0.086	1.9866	0.575
1 i 1	i i i i	4	0.005	0.005	1.9904	0.738
i 🚺 i	ן ום י	5	-0.042	-0.056	2.2316	0.816
r р i		6	0.119	0.109	4.2035	0.649
1 0 1	ן ון ו	7	-0.058	-0.043	4.6857	0.698
1 0 1	ן ום י	8	-0.039	-0.056	4.9010	0.768
· 🛑 ·		9	0.066	0.055	5.5295	0.786
– –		10	-0.168	-0.156	9.6376	0.473
1 1	I I I I	11	0.013	0.011	9.6625	0.561
i 🏚 i	ן ום ו	12	0.070	0.078	10.394	0.581
1 1	ן ום ו	13	0.022	0.056	10.463	0.656
1 1	1 1 1 1	14	0.006	0.010	10.469	0.727
· 📮 ·	ן וויין	15	0.084	0.038	11.524	0.715
1 0 1	1 1 1 1	16	-0.053	-0.019	11.950	0.747
· 🗖 ·	י 🗖 י	17	-0.085	-0.116	13.049	0.733
1 1 1	1 1	18	0.020	-0.017	13.112	0.785
1 0 1	1 1 1	19	-0.061	-0.024	13.691	0.801
	ינקי	20	-0.045	-0.054	14.007	0.830
· 🛛 ·	י מי	21	0.059	0.057	14.565	0.844
	1 1 1 1	22	-0.031	0.017	14.717	0.874
1 1		23	0.007	0.040	14.726	0.904
ا ا ا	יםי	24	-0.068	-0.103	15.480	0.906
1 0 1	ן יםי	25	-0.088	-0.090	16.763	0.890
1 D 1		26	0.068	0.086	17.539	0.892
1] 1	1 1	27	0.013	-0.023	17.566	0.916
1	 	28	-0.029	-0.018	17.713	0.933
		29	-0.041	-0.053	18.009	0.944
· I ·	ים י	30	-0.111	-0.128	20.144	0.913
		31	-0.038	0.017	20.392	0.927
u '	""	32	-0.140	-0.144	23.867	0.849
· Щ ·		33	-0.059	-0.047	24.500	0.857
· U ·	יוןי	34	-0.048	-0.050	24.921	0.872
· U ·		35	0.040	0.036	25.209	0.889
·Щ ·	<u> </u>	36	-0.072	-0.009	26.153	0.886

Table C.5: Δ SWAP2Y = Φ^2 (C, Δ TB3M, Δ COREPCE, Δ IPYOY, Δ LNSP500, Δ LNEURO, Δ LNVIX)

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob*
		1 0.031	0.031	0.1333	0.715
1 i 1	j i j i	2 -0.021	-0.022	0.1951	0.907
ı İ BI	j i 🗖 i	3 0.142	0.144	2.9699	0.396
i 🖬 i	ı ⊟ ı	4 -0.096	-0.108	4.2392	0.375
1 1 1		5 0.011	0.028	4.2571	0.513
r 🍅 i		6 0.135	0.111	6.8245	0.337
1 0 1	ן ון ו	7 -0.045	-0.028	7.1067	0.418
1 ()	ן ון ו	8 -0.022	-0.029	7.1778	0.518
1 0 1	ים, י	9 -0.051	-0.085	7.5579	0.579
י ם י	וםי	10 -0.097	-0.061	8.9328	0.538
i 📮 i	ן וים	11 0.067	0.071	9.5826	0.568
i 📮 i	ן וים	12 0.080	0.074	10.529	0.570
1 1 1	1 1 1	13 -0.035	-0.023	10.709	0.635
1 D 1	ן וףי	14 0.068	0.049	11.393	0.655
· 📮 ·	ı -	15 0.136	0.149	14.203	0.510
· 🗖 ·	ים י	16 -0.089	-0.071	15.404	0.495
. 	ים י	17 -0.042	-0.087	15.673	0.547
· 🛛 ·	1 1 1	18 0.067	0.024	16.372	0.567
· ·		19 -0.191	-0.165	22.098	0.279
1 1	1 1	20 -0.008	-0.001	22.108	0.335
1 Q 1	1 1 1 1	21 0.038	0.008	22.344	0.380
· ا ب	1 1 1 1	22 -0.086	-0.004	23.522	0.373
1 Q 1	1 1 1 1	23 0.028	0.023	23.646	0.424
ים י	ים, י	24 -0.093	-0.092	25.067	0.402
	יםי ו	25 -0.162	-0.108	29.428	0.246
1 1	יםי ו	26 -0.017	-0.084	29.476	0.290
10	ן ינן י	27 -0.049	-0.048	29.880	0.320
1 Q 1	1 1	28 -0.032	-0.009	30.055	0.361
1 1 1	וויי	29 0.018	-0.042	30.113	0.408
· 🗓 ·	1 1 1 1	30 -0.051	-0.008	30.561	0.437
– •	ים,	31 -0.179	-0.125	36.188	0.239
· Щ ·	ים, י	32 -0.076	-0.083	37.202	0.242
1 🚺 1		33 -0.015	-0.025	37.245	0.280
۲ Щ - ۲	ן יםי	34 -0.077	-0.055	38.318	0.280
· 📮 ·	ļ ' ∎'	35 0.142	0.125	41.972	0.194
· •	I <mark> </mark> I	36 -0.004	-0.030	41.975	0.228

Table C.6: Δ SWAP5Y = $\Phi^2(C, \Delta$ TB3M, Δ COREPCE, Δ IPYOY, Δ LNSP500, Δ LNEURO, Δ LNVIX)

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
.) .		1	0.029	0.029	0.1125	0.737
	j (j)	2	-0.040	-0.041	0.3300	0.848
· 🗖 ·	j j <u>p</u> i	3	0.105	0.108	1.8515	0.604
ı 🗖 i	ו בי ד	4	-0.086	-0.095	2.8630	0.581
i 🏚 i) (þ .	5	0.030	0.047	2.9854	0.702
· 🗐 ·		6	0.105	0.084	4.5371	0.604
1 0 1	ן וםי	7	-0.090	-0.078	5.6811	0.577
1 0 1	ים,	8	-0.066	-0.068	6.3046	0.613
10	ים,	9	-0.055	-0.072	6.7456	0.664
1 0 1		10	-0.053	-0.021	7.1500	0.711
i 🏚 i	ı p ı	11	0.097	0.092	8.5320	0.665
1 D 1		12	0.047	0.037	8.8645	0.714
ا ا	ים, י	13	-0.104	-0.087	10.463	0.656
i Di i	ן וני	14	0.058	0.057	10.973	0.688
· 🗐 ·	(p)	15	0.123	0.130	13.256	0.583
יםי	ים י	16	-0.078	-0.078	14.173	0.586
1 1	ן ון ו	17	0.011	-0.041	14.192	0.653
· 🕛 ·	ן ו	18	0.053	0.042	14.621	0.688
· ·	יםי ו	19	-0.177	-0.129	19.502	0.425
1 1	1 1 1 1	20	0.011	0.003	19.522	0.488
1 1	l (l)	21	0.008	-0.022	19.532	0.551
יםי	ļ (ļ)	22	-0.071	-0.018	20.341	0.562
1 1	I]I	23	0.013	-0.006	20.371	0.619
	יייי	24	-0.043	-0.014	20.669	0.658
1	ים י	25	-0.084	-0.058	21.843	0.645
	' U '	26	-0.042	-0.111	22.138	0.681
ישי	ישי	27	-0.074	-0.062	23.057	0.682
· II ·	'!'	28	-0.042	-0.024	23.363	0.715
·Ш·	' ! '	29	0.067	0.024	24.123	0.723
		30	-0.047	-0.036	24.503	0.749
		31	-0.140	-0.115	27.916	0.626
· • •		32	-0.033	-0.073	28.104	0.664
1 1 1		33	-0.051	-0.042	28.563	0.688
		34	-0.056	-0.037	29.135	0.705
		35	0.161	0.109	33.852	0.523
<u> </u>		36	0.016	0.018	33.898	0.569

Table C.7: Δ SWAP10Y = Φ^2 (C, Δ TB3M, Δ COREPCE, Δ IPYOY, Δ LNSP500, Δ LNEURO, Δ LNVIX)

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob*
		1	-0.009	-0.009	0.0103	0.919
. () .	ן ון ו	İ 2	-0.030	-0.030	0.1312	0.937
ı İ ı	j j	3	0.053	0.053	0.5172	0.915
1 0 1	i i 🖬 i	4	-0.098	-0.099	1.8497	0.763
i 🏮 i		5	0.027	0.029	1.9503	0.856
i 🏚 i		6	0.078	0.070	2.7995	0.834
1 0 1	ים,	7	-0.077	-0.066	3.6428	0.820
1 🛄 1	ן ים,	8	-0.070	-0.080	4.3489	0.824
1 0 1	ן ים,	9	-0.048	-0.055	4.6744	0.862
1 🛄 1	וםי	10	-0.072	-0.057	5.4228	0.861
· 🗖 ·		11	0.106	0.095	7.0516	0.795
i 🏚 i	ן וףי	12	0.056	0.045	7.5166	0.822
ا ا	l ı ∎ı	13	-0.126	-0.116	9.8910	0.703
1 () 1	ן וויין	14	0.039	0.029	10.123	0.753
· 📮 ·	ים ו	15	0.086	0.102	11.241	0.735
· 🛛 ·	יםי	16	-0.069	-0.060	11.969	0.746
1 () 1	111	17	0.041	-0.016	12.226	0.786
1 1 1	1 1	18	0.008	0.005	12.235	0.835
· ا ا	יםי	19	-0.107	-0.060	14.034	0.782
1 1 1	I I	20	-0.003	-0.026	14.035	0.829
1	וויי	21	-0.021	-0.026	14.107	0.865
1	111	22	-0.024	-0.013	14.197	0.894
1 [1	ן ים, י	23	-0.019	-0.052	14.257	0.919
1] 1	I]I	24	0.004	0.039	14.261	0.941
1 Q 1	1 1	25	-0.035	-0.020	14.462	0.953
1	ן ום י	26	-0.081	-0.139	15.548	0.946
1	ווייי	27	-0.042	-0.049	15.842	0.956
· •		28	-0.025	-0.006	15.947	0.967
		29	0.097	0.067	17.549	0.953
	יםי	30	-0.070	-0.094	18.387	0.952
· U	ישי	31	-0.077	-0.070	19.416	0.948
' <u></u> '		32	0.008	-0.015	19.428	0.960
' U '	'¶'	33	-0.105	-0.098	21.382	0.941
	ן י ע י	34	-0.018	-0.037	21.440	0.954
	l ' ₽ '	35	0.156	0.111	25.891	0.868
· •		36	0.009	0.014	25.905	0.893

Table C.8: Δ SWAP30Y = Φ^2 (C, Δ TB3M, Δ COREPCE, Δ IPYOY, Δ LNSP500, Δ LNEURO, Δ LNVIX)

	∆SWAP2Y	∆SWAP2Y	∆SWAP5Y	∆SWAP5Y	∆SWAP10Y	∆SWAP10Y	∆SWAP30Y	∆SWAP30Y
Main Equation								
ΔTB6M	0.73***	0.61***	0.62***	0.40***	0.54***	0.31***	0.56***	0.27***
	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.01)
Δ TB6M(-1)	-0.11	, , ,		<u>``</u>	, , ,		-0.22***	, , ,
	(0.15)						(0.00)	
Δ SWAP_Y(-1)				0.23***				
				(0.00)				
Δ SWAP_Y(-2)				-0.03				
				(0.59)				
Δ SWAP_Y(-3)				0.17**				
				(0.01)				
Δ SWAP_Y(-4)				-0.12*				
				(0.07)				
∆CORECPI	0.05	0.06	-0.007	0.01	-0.08	-0.04	-0.11	-0.09
	(0.23)	(0.20)	(0.90)	(0.85)	(0.20)	(0.37)	(0.05)	(0.09)
ΔΙΡΥΟΥ	-0.004	-0.006	0.0001	-0.004	0.001	-0.0005	0.004	-0.002
	(0.26)	(0.11)	(0.97)	(0.30)	(0.64)	(0.90)	(0.23)	(0.72)
ΔLNDJIA		0.69		1.60		1.49		1.55
		(0.04)		(0.00)		(0.02)		(0.02)
ΔLNYEM		1.39		2.59		2.79		2.12*
		(0.00)		(0.00)		(0.00)		(0.00)
ΔLNVIX		0.03		0.12		0.05		-0.01
		(0.54)		(0.21)		(0.68)		(0.94)
Intercept	-0.001	-0.008	-0.005	-0.02	-0.01	-0.03	-0.01	-0.03
	(0.83)	(0.31)	(0.71)	(0.05)	(0.58)	(0.03)	(0.50)	(0.02)
			Model I	nformation				
Obs	133	133	133	133	133	133	133	133
Adj R ²	0.62	0.68	0.30	0.50	0.23	0.41	0.27	0.44
AIC	-2.06	-2.21	- 1.13	- 1.44	-0.89	-1.14	-0.95	– 1.19
T I CI IN	44.20	41.00	Diagno	ostic Tests	10.04	14.24	10 75	15.60
Joint Significance	44.30	41.09	15.36	14.30	10.84	14.24	10.75	15.62
F-Test	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)	(0.00)
Serial Correlation	1.97	1.80	1.88	1.85	1.89	1.88	1.96	1.84
Durbin-Watson								
Stat Sorial Correlation	0.27	0.70	0.46	0.80	0.67	0.76	0.67	1 49
Brougeh Codfroy	(0.27)	0.70	(0.40)	(0.45)	(0.51)	(0.70)	(0.51)	1.48
I M Tost	(0.76)	(0.50)	(0.05)	(0.43)	(0.31)	(0.47)	(0.31)	(0.23)
Hotoroskodosticity	0.01	1.68	0.48	2.00	0.27	3 71	0.17	5.13
Brousch-Pagan-	(0.91)	(0.12)	(0.48)	(0.03)	(0.27)	(0,00)	(0.07)	(0,00)
Godfrey Test	(0.40)	(0.12)	(0.73)	(0.03)	(0.09)	(0.00)	(0.77)	(0.00)
Normality Test	1.12	0.96	7.65	21.76	15 53	14 44	27.70	1 18
Jarque-Bera Stat	(0.57)	(0.62)	(0.02)	(0.00)	(0.00)	(0.00)	(0.00)	(0.55)
Stability	7.85	12.82	7.85	5.34	0.70	3.57	0.51	2.10
Diagnostic	(0.00)	(0.00)	(0.00)	(0.01)	(0.50)	(0.03)	(0.60)	(0.13)
Ramsey RESET	(0.00)	(0.00)	(0.00)	(0.01)	(0.00)	(0.00)	(0.00)	(0.10)
Test								
	1							1

Appendix D: Additional Regressions: ARDL (p, q) Model (with Δ TB6M and a set of alternative control variables)

Note: all vars are in diff, *p*-values are in parenthesis. BG LM and Ramsey RESET tests are with 2 lags.

Models	∆SWAP2Y	∆SWAP5Y	∆SWAP10Y	∆SWAP30Y				
Lags	Panel One							
1	9.33	0.68	1.80	2.39				
	(0.00)	(0.41)	(0.18)	(0.12)				
4	3.67	0.46	1.26	1.95				
	(0.01)	(0.76)	(0.29)	(0.32)				
8	2.69	0.92	1.03	0.48				
	(0.01)	(0.50)	(0.41)	(0.87)				
12	2.15	0.82	0.77	0.41				
	(0.02)	(0.63)	(0.68)	(0.96)				
Lags		Panel Two						
1	9.56	0.07	0.003	0.09				
	(0.00)	(0.79)	(0.95)	(0.77)				
4	3.92	0.14	0.35	0.30				
	(0.00)	(0.97)	(0.84)	(0.88)				
8	2.43	0.54	0.82	0.50				
	(0.02)	(0.82)	(0.59)	(0.85)				
12	2.11	0.62	0.74	0.39				
	(0.02)	(0.82)	(0.71)	(0.96)				

Appendix E: ARCH LM Test with ARDL (p, q) Models (with \triangle TB3M)

Note: ARDL (p, q) models (from table 4) include the change in the short-term interest rate (Δ TB3M) and the controls (namely Δ COREPCE and Δ IPYOY in panel one for the basic model and Δ COREPCE, Δ IPYOY, Δ LNSP500, Δ LNEURO, and Δ LNVIX in panel two for the extended model); *p*-values are in parenthesis.