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Intergenerational scars: The impact of parental unemployment on individual health later in life*

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Abstract

This paper studies whether individuals that experienced parental unemployment during their childhood/early adolescence have poorer health once they reach the adulthood. We used data from the German Socio-Economic Panel from 2002 until 2018. Our identification strategy of the causal effect of parental unemployment relied on plant closures as exogenous variation of the individual labor market condition. We combined matching methods and parametric estimation to strengthen the causal interpretation of the estimates. On the one hand, we found a nil effect for parental unemployment on mental health. On the other hand, we detected a negative effect on physical health. The latter is stronger if parental unemployment occurred in early periods of the childhood, and it is heterogeneous across gender. The negative effect of parental unemployment on physical health may be explained by a higher alcohol and tobacco consumption later in life.

Keywords: Parental unemployment; plant closure; mental health; physical health; health behaviors

JEL Classification: I14, J13, J62, J65

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1 Introduction

Labor economists, psychologists and social scientists have extensively studied the consequences of unemployment on the lives of the individuals and of their relatives ([Jahoda, 1982](#)), especially to understand whether unemployment affects both economic and non-economic outcomes ([Janlert and Hammarström, 2009](#)) and how long these effects last.

Past empirical research provided evidence about the presence of scarring effects on wages, (re-)employment probability, mental health and overall life satisfaction (see e.g. [Jacobson et al., 1993](#); [Winkelmann and Winkelmann, 1998](#); [Arulampalam, 2001](#); [Chan and Stevens, 2001](#); [Kassenboehmer and Haisken-DeNew, 2009](#); [Green, 2011](#); [Drydakis, 2015](#)). For the labor market, the findings were remarkably consistent, regardless of the period and country under analysis, the type of data used and the identification strategy employed ([Filomena, 2023](#)). [Picchio and Ubaldi \(2022\)](#) provided an up-to-date meta-analysis on the health effects of unemployment and pointed out a negative effect for laid-off workers, especially on mental health and life satisfaction. The effect is less severe for the relatives. In addition, their heterogeneity analysis revealed the need to control for liquidity constraints or the reason why someone lost the job.

A branch of the literature put special emphasis on the relationship between parents and children, and studied whether and to what extent the unemployment of the former may shape the outcomes of the latter. This relationship may be one of the most important in social sciences, because it may be the principal channel through which the phenomenon of intergenerational mobility takes place ([Becker and Tomes, 1979, 1986](#)). Past empirical research on parental unemployment mostly focused on children's educational and labor market outcomes. Whilst there is consensus on the negative effects on school ambitions and performances, grade completion and tertiary education enrollment (see e.g. [Kalil and Ziol-Guest, 2008](#); [Coelli, 2011](#); [Rege et al., 2011](#); [Stevens and Schaller, 2011](#); [Mörk et al., 2020](#)), the findings for the labor market outcomes instead present some degree of heterogeneity. In the US, [Oreopoulos et al. \(2008\)](#) showed that the exposure to the paternal job loss during the years of the adolescence leads the individuals to earn less and rely more on welfare and social aids once they reach the adulthood. For the UK, [Macmillan \(2014\)](#) found that paternal unemployment is also associated with an increase in the duration of the potential joblessness spells experienced by these children. The effect was worse in areas with precarious labor market conditions. Using an instrumental variables approach, [Hérault and Kalb \(2016\)](#) confirmed these findings also in Australia. Finally,

individuals that suffered parental unemployment also achieved worst socioeconomic status once they grew old (Karhula et al., 2017). Nevertheless, Bratberg et al. (2008) and Müller et al. (2017) failed to find similar findings in Norway and Germany. Ekhaugen (2009) posited that most of the negative effect found is the reflection of a simple intergenerational correlation which disappeared once the unobserved heterogeneity at family level was appropriately controlled for. In this regard, Ekhaugen (2009) compared achievements from siblings pairs once they reached the adulthood and found that the net effect was not statistically different from zero and, if anything, positive.

This paper is on the impact of parental unemployment on children's health later in life. In the literature about the consequences of parental unemployment, only few studies tackled the health perspective, with the majority focusing on short-term effects while neglecting possible long-run repercussions (see e.g. Lindo, 2011; Powdthavee and Vernoit, 2013; Pieters and Rawlings, 2020). The purpose of this paper was to shed light on the long-term effects by answering the following research questions: i) Does parental unemployment exert a long-term effect on children's health? ii) If yes, which health dimension is impaired? iii) Finally, at what extent?

We used data from the German Socio-Economic Panel (SOEP) for the period 2002 to 2018. The SOEP is a rich longitudinal dataset, representative of the adult German population. It contains a wide range of demographic and socio-economic information about the individual and the household. The survey started in 1984 and takes place annually ever since. We used the Mental Component Summary (MCS) and Physical Component Summary (PCS) scales as outcomes variables (Ware et al., 2000). These measures are widely employed in the economic and psychological literature and nowadays represent a standard in those studies seeking to measure the health effects of unemployment (Schmitz, 2011; Marcus, 2013; Neubert et al., 2019; Stauder, 2019). They are provided by the SOEP version of the Short Form Health Survey (SF-12v2) questionnaire, which is the shorter and more practical version of the 36-items Short-Form Health Survey (SF-36). The SF-12v2 is one of the most popular tools used to collect information about individual health and quality of life in large survey samples. We observed the health outcomes when the individuals were adults (i.e. when they were 18 to 31 years old).

We built the identification strategy upon two stepping stones. First, we exploited the information about the reason for unemployment entry to tackle the endogeneity problem typical of the relation between unemployment and health (Kassenboehmer and Haiken-DeNew, 2009). We selected into the treated group only those children whose parent

experienced unemployment due to a plant closure. A plant closure is widely regarded to be a quasi-experimental variation of the individual labor market condition, uncorrelated with the unobserved characteristics of the laid-off worker (Brand, 2015). We defined parental unemployment using three criteria: i) at least one parent lost his/her job due to a plant closure; ii) s/he officially registered the unemployment entry at the German Employment Agency; iii) the layoff occurred when the individual was between 0 and 15 years old. Second, we applied a double-robust approach that combines matching methods and standard parametric estimation in order to reduce the risk of model dependence and hence add further credibility to the causal interpretation of our estimates (Ho et al., 2007).

The main contribution of this paper is to study the long-term effect of parental unemployment on health by a methodological approach which combines an exogenous variation of the individual labor market condition with estimation techniques which are robust to model dependence. To the best of our knowledge, there are other two papers which tried to shed light on the potential long-term effects of parental unemployment on individuals' health, namely Lam and Ambrey (2019) and Nikolova and Nikolaev (2021). However, our study differed under some important aspects. Compared to the former, we applied a narrower definition of unemployment, which is based on plant closures as exogenous reason for entering unemployment. Moreover, we collected the information on parental unemployment directly from the parents' files, reducing the risk of reporting bias. Compared to the latter, who dealt with life satisfaction as dependent variable, we dug more deeply by focusing on specific measures of mental and physical health. In fact, even though the correlation between life satisfaction and mental health is generally high, they are distinct phenomena (Headey et al., 1993). Moreover, we enlarge the picture to measures of physical health and health behaviors, like drinking alcohol and smoking. Finally, we applied bias-adjusted methods robust to parametric misspecification to strengthen the causal interpretation our findings.

We found mixed effects of the parental unemployment on individuals' health later in life: a negative effect on physical health and a nil effect on mental health. About physical health, the effect is stronger if parental unemployment occurred during the early childhood and heterogeneous across gender. Maternal unemployment matters if it occurred during the early childhood, whereas paternal unemployment matters if it was experienced later. With respect to the gender of the kid, daughters were negatively affected, whilst sons were not. We also found that parental unemployment induced a higher probability of alcohol and tobacco consumption later in life.

This article is organized as follows. Section 2 presents both theoretical predictions and empirical evidence on the effects of parental unemployment on individual outcomes. Section 3 presents the dataset, the sample selection criteria, the variables and the identification strategy. Section 4 reports and comments on the findings. Section 5 concludes.

2 Literature review

The economic literature about intergenerational transmission and child development started with the seminal work of [Becker and Tomes \(1979, 1986\)](#). The Beckerian model assumed that individuals' lives may be divided in two periods: childhood and adulthood. The first period is characterized by individuals receiving and transforming parental investments in marketable skills and human capital. The second period is characterized by individuals generating returns by exploiting these skills on the labor market ([Becker and Tomes, 1979, 1986](#)). The implicit assumption of the model is that the parental investments are perfect substitutes in the individual human capital production function.

[Cunha and Heckman \(2007\)](#) criticized this idea, arguing that parental investments are not equally effective at any point in time of the childhood. Childhood is a complex and heterogeneous moment in the lives of the individuals. It is composed by many different stages, with different impacts in the human development. A multistage framework is to be preferred over a one period approach, with each stage constituting the base of the next one, so as to account for the complex dynamics behind the human capital formation process and to put the necessary emphasis on those 'sensitive periods' characterizing child development. Because parental inputs are transformed in human capital and since human capital tends to both accumulate over time and increase in efficiency, investments made in the early periods of the childhood may generate higher returns compared to those made later. In the same spirit, setbacks suffered in the early periods of the childhood may generate greater losses in the future.

Although parental unemployment may play a relevant role in the human capital formation process, assessing the sign of the relationship is not trivial because parental unemployment could exert both positive and negative effects. On the one hand, positive effects may arise because unemployment increases the amount of leisure time available for the laid-off parent ([Knabe et al., 2010](#)). The laid-off parent may decide to use this additional time for a more intensive parenting activity. Parental involvement is an essential input in the human capital production function. Through the parenting style, a parent can

shape his/her child's preferences (Doepke et al., 2019). Parental involvement also prevents the child to achieve sub-optimal outcomes later in life (Ermisch and Francesconi, 2013). Moreover, it promotes the development of soft skills like resilience, which further enhance the stock of human capital (Masten and Narayan, 2012; Masten and Palmer, 2019). On the other hand, negative effects might arise because families risk to end up in economic and emotional deprivation (Jahoda, 1982; Janlert and Hammarström, 2009). Unemployment reduces the economic resources available (Jacobson et al., 1993; Arulampalam, 2001). This reduction may offset (or at least partially hinder) the parental ability to invest in educational and health goods for the child (Grossman, 1972; Becker and Tomes, 1979, 1986). Unemployment also causes emotional distress, that in turn may lead to family breaks (Jensen and Smith, 1990; Eliason, 2012; Marcus, 2013; Mörk et al., 2020; Di Nallo et al., 2022). Disrupted family structures represent inhospitable environments for the child development. For instance, in the UK Ermisch and Francesconi (2001) and Ermisch et al. (2004) showed that children that grew up in separated or single-parent families reported poorer educational, behavioral and health outcomes later in life. Francesconi et al. (2010) reported similar findings in Germany. Moreover, unemployment increases the likelihood for the laid-off worker of engaging in risky behaviors, such as smoking or drinking (Schunck and Rogge, 2010; Reine et al., 2013; Mörk et al., 2020). Past empirical research showed that these behaviors are likely to transmit from parents to children, representing another potential threat in children's human capital accumulation process (Bantle and Haisken-DeNew, 2002; Yu, 2003; Göhlmann et al., 2010; Schmidt and Tauchmann, 2011).

Also the handful of studies dealing with the health consequences of parental unemployment leaned more towards the negative effects, albeit without providing a clear-cut picture. In Sweden, Mörk et al. (2020) found that parental joblessness did not impair the mental health of the offspring in the childhood. However, in the US and Australia, Brand and Thomas (2014) and Lam and Ambrey (2019) found that it led to more depressive symptoms in the adulthood. Nikolova and Nikolaev (2021) realised that in Germany such negative effect emerged only if parental unemployment occurred either at early childhood or at the early adolescence (i.e. 0-5 and 11-15). Powdthavee and Vernoit (2013) found mixed evidence on the happiness of the British adolescents, with paternal unemployment exerting a positive effect and maternal unemployment a negative one. Haisken-DeNew and Kind (2012) reported similar findings on the life satisfaction of young adults in Germany. Pieters and Rawlings (2020) found that in China, whereas paternal joblessness was

associated with poorer dietary habits and increased odds of low-weight births, maternal unemployment generated an improvement in the dietary habits. [Lindo \(2011\)](#) reported similar findings on birth weight in Germany. In particular, the author found that those children born after a paternal job displacement were more likely to report a 5% lower weight. Finally, in Ireland, [Briody \(2021\)](#) reported that children living in a household with at least one unemployed parent were more likely to consume unhealthy food and exercise less. These associations were particularly strong in the case of paternal unemployment.

3 Data and Method

3.1 Data and sample selection criteria

We obtained our sample by combining two datasets: i) the German Socio-Economic Panel (SOEP-Core v36), a household panel survey representative of the adult German population ([Wagner et al., 2007](#)); ii) the German Employment Agency statistics. The SOEP represents the main building block. Started in 1984, the survey is carried out annually. Approximately 20,000 adult individuals join the panel every year.¹ The SOEP presents several useful features. First, once a household is selected, all members are interviewed and they may remain in the sample even if they separate from the original household. If an individual gets married or starts a cohabitation, also the partner and the future kids join the sample.² This longitudinal structure allows to track individuals over time, regardless of complex family dynamics. Individuals exit the sample only if they die, move abroad or decide to quit.³ Second, the SOEP collects retrospective information. By combining both retrospective and currently acquired information, a researcher can recreate the full history of each SOEP member. Third, the SOEP covers a vast range of topics, such as health and consumption habits, well-being, labor conditions, political affiliations, income, etc.

We started with 149,565 individuals for a total of 1,164,296 observations. We kept only those individuals with both parents present in the sample. We required these parents

¹ According to the SOEP, a person is considered adult only after s/he turned 17.

² Either one of the parents or a legal guardian answers the survey questions on behalf of non-adults.

³ In the SOEP, panel attrition is not a major problem. The percentage of successful follow-ups is above 95% in all years. Furthermore, participation rates are usually beyond 85% ([Siegers et al., 2021](#)). The SOEP is continuously integrated with refreshment samples to maintain the representativeness of the German population ([Goebel et al., 2019](#)).

to be married or cohabiting. We restricted on individuals born from 1984 onward and aged between 18 to 31 years old when interviewed. In addition, we required individuals to have lived with their parents until they turned 15. Then, we dropped individuals who had missing observations in the outcome variables. Finally, we removed those individuals with at least one parent who experienced unemployment not due to plant closure. We ended up with 2,026 individuals for a total of 5,161 observations. Table 1 reports how the sample size shrank through our selection process.

Table 1: Sample size across the selection process

	Individuals in the sample	Individual-per-year observations (<i>n</i>)	Dropped observations
Initial gross merged sample	149,565	1,164,296	–
After removing individuals with no parents identifiers	21,512	774,432	389,864
After keeping individuals with both parents identifiers [†]	17,109	615,924	158,508
After removing individuals born before 1984	9,861	354,996	260,928
After keeping individuals aged 18 to 31	9,220	78,030	276,966
After keeping individuals who lived with their parents until 15 years old	6,149	30,093	47,937
After removing missing values in the outcome variables	4,858	11,424	18,669
After removing individuals with at least one parent having experienced unemployment not due to plant closure	2,026	5,161	6,263
Final sample	2,026	5,161	

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

[†] Parents must be married or cohabiting.

We applied these sample selection criteria for different reasons. First, by requiring parents to be in the sample, we are able to collect information on parental unemployment directly from their files, with their children providing information about themselves when adult. In this way, we reduced the risk of reporting bias, i.e. the measurement error due to an individual misreporting information, in this case, about someone else. Second, by requiring parents to be married or cohabiting, we excluded disrupted family environments and isolated the effect of the parental unemployment from other potential relevant channels, like parental divorce or growing up in a single-parent family. Third, by focusing on children born only from 1984, we can observe the full history of each individual and his/her family and we avoided losing potentially relevant information. Fourth, by requiring individuals to have spent their whole childhood living with their parents, we are sure that the kid was fully exposed to parental unemployment when it occurred. Finally, with the last sample selection criterion, we leave in our sample only individuals who either did not experience parental unemployment during their childhood (the control units) or experienced parental unemployment due to plant closure during their childhood (the treated

units). The latter sample selection criterion is the pillar of our identification strategy and is discussed in the next subsection.

3.2 Identification strategy

Identifying the causal effect of parental unemployment on children’s health later in life may be tricky due to omitted variables ([Kassenboehmer and Haisken-DeNew, 2009](#); [Brand, 2015](#); [Farré et al., 2018](#)). In our framework, omitted variables are those unobservables which are jointly correlated with the likelihood of parental unemployment and children’s health later in life, for example family and social background, predetermined and innate offspring’s health conditions and household wealth. We tackled this endogeneity concern by using plant closures as the reason for unemployment entry. We assumed that plant closures are uncorrelated with the unobserved characteristics of the laid-off worker. Plant closures are widely regarded as a quasi-experimental variation of the individual labor market condition ([Brand, 2015](#)). However, a plant closure rarely happens overnight. Typically a public announcement is made in advance. The time interval between the announcement and the shutdown may represent a ‘breeding ground’ for a selectivity bias to arise. Those workers who are more likely to find a new job elsewhere may anticipate the plant closure, look for a new job before the plant closure and avoid the transition into unemployment. Those who do not anticipate the plant closure and/or are not able to find a new job before the layoff experience transit to unemployment. Therefore, they may not be a random sample from the workforce and they may be endowed with labor market characteristics systematically different from those of individuals at work. We tried to limit these concerns by including in our empirical model a rich set of control variables predetermined with respect to the treatment (e.g. parental migration background, parental education, parental cumulative unemployment and parental job loss expectations) and by checking whether the estimated parental unemployment effect changed across the various model specifications ([Marcus, 2013](#); [Nikolova and Ayhan, 2019](#); [Nikolova and Nikolaev, 2021](#)).

We observed the treatment and the outcomes in two different moments. The outcome variables were observed when the individuals were adults (i.e. between 18 and 31 years of age) whereas the treatment was observed when they were children (i.e. 0 to 15 years old). We followed [Cunha and Heckman \(2007\)](#) and tested whether parental unemployment exerts different effects on human capital of the individuals depending on the timing of its

occurrence. We split the childhood in three periods: 0-5 years of age for early childhood, 6-10 for late childhood and 11-15 for early adolescence (Ermisch and Francesconi, 2001; Ermisch et al., 2004). Hence, we estimated three different models, one for each childhood period.

For individual i at time t , the health outcome variable y_{it} is specified as follows:

$$y_{it} = \delta U_i + \beta \mathbf{X}_{it} + \alpha_i + \varepsilon_{it}, \quad (1)$$

where U_i is the binary indicator for parental unemployment, \mathbf{X}_{it} is a vector of individual and family controls and α_i and ε_{it} are the error term components.⁴ Because parental unemployment is time invariant with respect to the outcome, we could not include fixed-effects; α_i is therefore assumed to be a random effect, uncorrelated with the covariates and with the treatment. Equation (1) imposes a particular parametric structure to the conditional mean of the dependent variable. If this is misspecified, the estimation of Equation (1), for instance by Generalized least Squares (GLS), may lead to a biased estimate of the treatment effect if the functional relationship with the control variables is misspecified or if there are areas outside of common support which require extrapolation and therefore more model dependent (Ho et al., 2007; Peter, 2016). In order to avoid biases due to model dependence, we followed Ho et al. (2007) and applied their double-robust approach, which combines matching methods and standard parametric estimation techniques. More in detail, it consists in preprocessing the data over a set of observed confounders using matching methods so that the relationship between U_i and \mathbf{X}_{it} is eliminated or reduced. Thereafter, standard parametric estimation techniques are applied on the preprocessed sample. The matching part, by covariate balancing, is meant to reduce the risk of model dependence in the subsequent parametric estimation step, whereas the parametric estimation is meant to take care of the potentially leftover sample imbalance. Ho et al. (2007) argued that this combined approach is more likely to yield unbiased estimates under weak conditions: “if either the matching or the parametric model is correct, but not necessarily both, causal estimates will still be consistent”.

In the matching step, we combined two methods: i) the coarsened exact matching (CEM) (Iacus et al., 2008, 2011) and ii) the entropy balancing matching (EB) (Hainmueller, 2012). Hainmueller and Xu (2013) suggested that the combination of these two

⁴Although \mathbf{X}_{it} is indexed by t , all the covariates are either predetermined with respect to the treatment, and therefore constant over t , or they change over t following an exogenous path, like age.

methods make the subsequent estimates less exposed to the risk of model dependence, strengthening the causal interpretation.⁵ CEM reduces the general sample imbalance by producing an exact matching between the treated and the control observations. Those observations that do not find a counterfactual are discarded. Compared to other exact matching techniques, CEM suffers less from the curse of dimensionality, but it still remains highly sensitive to the number of matching variables included. The greater the number of the covariates, the larger is the proportion of observations discarded. Hence, CEM requires a minimal set of meaningful pre-treatment covariates. By discarding the non matched observations, the monotonic imbalance bounding property is respected by construction (Iacus et al., 2011). EB matching further increases the balance properties of the sample by reweighting the control units in a such way that the joint distribution satisfies some pre-specified moments conditions. The selected weights minimize an entropy distance metric over a r -dimensional set of constraints. We followed Marcus (2013) and imposed the conditions on the first and second moments of the full set of pre-treatment variables.

We started with the CEM and then applied the EB matching on the CEM-restricted sample. In the CEM, we matched our observations over a set of basic parental characteristics: parental migration background, parental educational level and parental age when the child was born. We then discarded those observations that were unmatched. This should have increased the internal validity of the subsequent estimates as we retained only individuals with similar parental background. In the EB, we enriched the set of pre-treatment characteristics and included the following variables: child gender, child age, child migration background, whether the child is the first-born, region where the child was born, paternal cumulative unemployment, maternal cumulative unemployment, average household income quintile during the treatment period and average number of siblings during the treatment period (Peter, 2016). Table 2 shows how the sample size changed due to the double matching algorithm.

In Appendix A, we reported the gains in balance for the mean and the variance of each covariate. The balance properties improved remarkably. In fact, the bias, i.e. the standardized difference between the treated and the matched controls, was reduced for almost every covariate. However, it came at the cost of losing observations, 65% on average across the three subsamples.

⁵We used the Stata commands `cem` and `ebalance` developed by Blackwell et al. (2009) and Hainmueller and Xu (2013), respectively.

Table 2: Sample size before and after double-matching

	Before matching		After matching	
	Individuals in the sample	Individual-per-year observations (n)	Individuals in the sample	Individual-per-year observations (n)
Controls and treated at 0 to 5	797	2,118	199	572
Controls and treated at 6 to 10	1,155	3,227	464	1,335
Controls and treated at 11 to 15	1,810	4,588	664	1,693

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Finally, we computed the Average Treatment Effect on the Treated (ATT) by estimating Equation (1) only on the matched observations and using the weights obtained from the EB. When the outcome is continuous, the ATT corresponds to the δ coefficient of the treatment indicator in the parametric specification (Jones et al., 2020). Since we used the matching methods only to reduce the covariates imbalance and not for estimation purposes, we could use standard methods for computing standard errors and making inference on the estimated coefficients (Ho et al., 2007; Jones et al., 2020). To account for within-individual correlation, we estimated standard errors using Liang and Zeger’s (1986) cluster-robust variance estimator.

3.3 Variables and descriptive statistics

3.3.1 Dependent variables

As health outcome variables, we selected the Mental Component Summary (MCS) and the Physical Component Summary (PCS) scales (Ware et al., 2000). These two indicators are built on the items available in the SOEP version of the Short Form Health Survey (SF-12v2) (Ware et al., 1996). The SOEP introduced this questionnaire for the first time in 2002 and ever since every two years. This survey covers several health dimensions like bodily pain, emotional draining, and social functioning, with each item referring to the last 4 weeks period of the life of the individual.⁶ Hence, the MCS and PCS are two general health indicators measuring the current health status.

We calculated the two indicators by a factor analysis. They take values from 0 to 100, with mean 50 and standard deviation 10 in the 2004 SOEP sample (see Andersen et al., 2007, for the full description).⁷ Higher scores indicate better health. The two variables

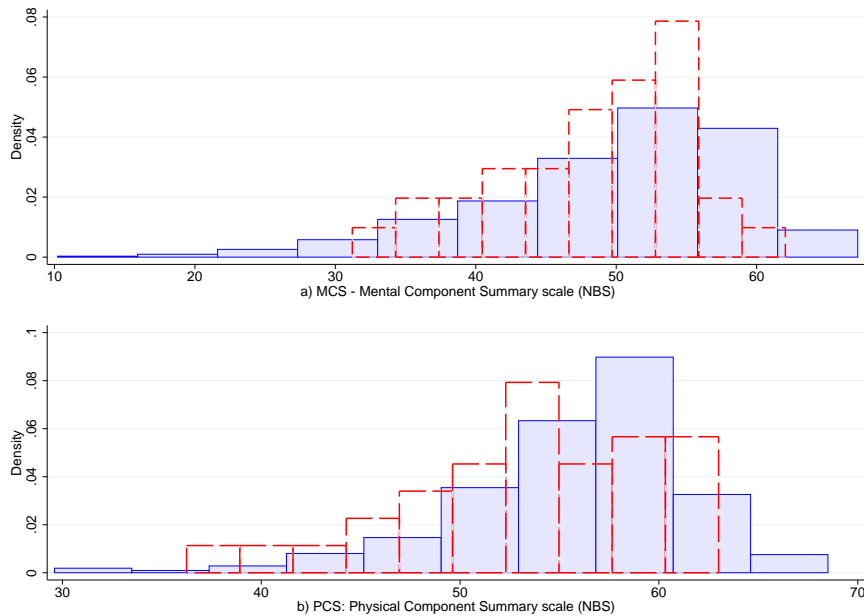
⁶Table B.1 in Appendix B presents a facsimile of the questionnaire.

⁷See also Nübling et al. (2006) for the open access SPSS code.

were proved to be valid and reliable health measures compared to other scales (Gill et al., 2007) and are widely used in the economic and psychological literature (Schmitz, 2011; Marcus, 2013; Neubert et al., 2019; Stauder, 2019).

Table 3 reports their summary statistics. Figures 1, 2 and 3 plot the density distributions of the outcome variables for each subsample and distinguishing between treated and control units. The treated individuals reported worse mental and physical health conditions on average if parental unemployment occurred when they were between 0 and 5 years old. Furthermore, the average physical health indicator is smaller for those individuals who were treated between 11 and 15 years of age. This descriptive evidence is similar to the one in Nikolova and Nikolaev (2021) for life satisfaction. Nikolova and Nikolaev (2021) suggested as a possible explication mechanism the presence of an underlying psychological trauma that was likely to accumulate over time. The chain of the events would have seen parental unemployment hitting first mental health and then affecting life satisfaction only later.

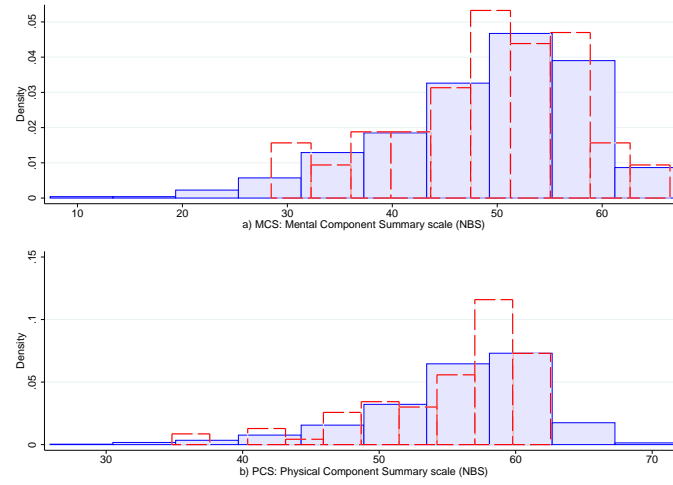
Figure 1: MCS and PCS density plots – Treatment at 0-5



Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: Panel a) refers to the Mental Component Summary (MCS) scale. Panel b) refers to the Physical Component Summary (PCS) scale. Colorized solid bins in blue refer to control units. Empty dashed bins in red refer to treated units. We used before-double matching observations.

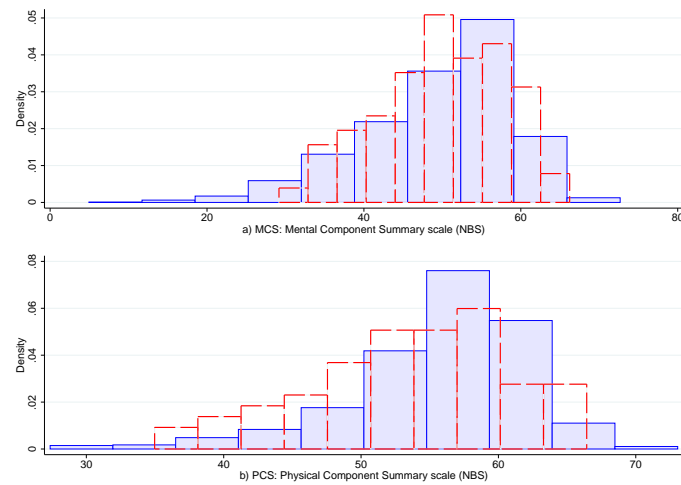
Figure 2: MCS and PCS density plots – Treatment at 6-10



Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: Panel a) refers to the Mental Component Summary (MCS) scale. Panel b) refers to the Physical Component Summary (PCS) scale. Colorized solid bins in blue refer to control units. Empty dashed bins in red refer to treated units. We used after-double matching observations.

Figure 3: MCS and PCS density plots – Treatment at 11-15



Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: Panel a) refers to the Mental Component Summary (MCS) scale. Panel b) refers to the Physical Component Summary (PCS) scale. Colorized solid bins in blue refer to control units. Empty dashed bins in red refer to treated units. We used after-double matching observations.

Table 3: Summary statistics of the MCS and PCS indicators

Variables	Treatment at	Treated					Controls				
		Obs.	Mean	σ	Min.	Max.	Obs.	Mean	σ	Min.	Max.
MCS	0-5	33	48.804	7.166	31.208	62.042	539	50.070	9.095	10.226	67.194
	6-10	84	49.100	8.653	28.459	66.461	1,251	49.010	9.477	7.394	67.194
	11-15	69	49.855	8.144	29.174	66.227	1,624	49.453	9.557	4.939	72.665
PCS	0-5	33	53.237	6.407	36.245	63.000	539	55.857	5.795	29.599	68.502
	6-10	84	55.161	5.977	34.821	62.554	1,251	55.778	6.465	27.364	73.053
	11-15	69	53.645	7.174	34.973	66.429	1,624	55.663	6.274	25.896	71.870

Source: German Socio-Economic Panel version 36 (SOEP-Core v36). We used after-double matching observations.

Table 4 reports the correlation matrix between our outcomes variables and the life satisfaction used by [Nikolova and Nikolaev \(2021\)](#). Despite a good level of correlation between mental health and life satisfaction, the figures in Table 3 suggest that previous findings on life satisfaction may be the consequence of both a physical and a psychological scar.⁸ On top of the speculation in [Nikolova and Nikolaev \(2021\)](#) for their findings, which was based on parental unemployment generating a psychological trauma, it may be that the treated individuals engaged in risky behaviors later in life, like smoking or drinking, which in turn impaired their physical health, leading to a decreased level of life satisfaction. One of the contributions of our analysis is that it shed light into the mediating channels possibly explaining the findings about life satisfaction in [Nikolova and Nikolaev \(2021\)](#).

Table 4: Correlation matrices between MCS, PCS and life satisfaction

	Treatment at 0-5			Treatment at 6-10			Treatment at 11-15		
	Life satisfaction	MCS	PCS	Life satisfaction	MCS	PCS	Life satisfaction	MCS	PCS
Life satisfaction	1.000			1.000			0.000		
MCS	0.383	1.000		0.390	1.000		0.432	1.000	
PCS	0.191	-0.086	1.000	0.189	-0.052	1.000	0.172	-0.059	1.000

Source: German Socio-Economic Panel version 36 (SOEP-Core v36). We used after-double matching observations.

3.3.2 Treatment variable

We defined the treatment using the information coming from two SOEP questions. The first question is about the reason for the job termination, which is part of the survey since 1985 and since 1991 it includes “due to shutdown” among the possible answers. The

⁸The correlation levels are line with other previous studies. For instance, using SOEP data, [Marcus \(2013\)](#) found that the ρ between mental health and life satisfaction was equal to 0.4.

second question is on the (un-)employment status. We defined as treated those individuals who, when they were between 0 and 15 years old, had at least one parent that was fired due to a plant shutdown and registered the unemployment event at the German Employment Agency. Individuals whose parents became unemployed for reasons different from a plant closure were excluded from the sample. We also excluded individuals whose parents were in other labor market status, like civil servants, retired or self-employed.⁹

For the sake of having similar control and treated units, we also required that the parents of both groups had the same initial labor market condition. Fathers had to be private employees before the shutdown, whilst mothers could be either private employees or not employed.¹⁰

3.3.3 Control variables

Following previous studies on parental unemployment effects on individuals' outcomes (see e.g. [Bratberg et al., 2008](#); [Oreopoulos et al., 2008](#); [Ekhaugen, 2009](#); [Siedler, 2011](#)), we included a large set of individual and family socio-economic characteristics. The individual characteristics were gender, age, migration background and whether the individual was the first-born child in his/her household. We also included the average number of siblings during the treatment period and three sets of dummy variables for the survey in which the individual was interviewed, the year of birth and the *länder* where the individual was born. This set of variables allowed to control for time and spatial heterogeneity. The family characteristics were the age that the parent had when the child was born, the parental migration background, the highest educational level achieved by the parent, the cumulative unemployment experience at the time of the treatment and the income quintile in which the household found itself during the treatment period ([Bratberg et al., 2008](#)). We included the household income variable in order to separate the potential poverty component from the parental unemployment effect ([Siedler, 2011](#)). Furthermore, we controlled for local labor market conditions at the time of the eventual treatment by using the aver-

⁹The definition of the treatment is based on the idea that a person may lose his/her own job due to plant closure. Neither the self-employed nor the retired can experience this event. For civil servants, the occurrence of the event is very unlikely, since they enjoy higher job protection than private employees.

¹⁰We defined non-working mothers using the SOEP item for the occupational status. We used the category "*Not employed (NE): without further information*". The condition on the mothers was meant to avoid additional possible sample selectivity ([Heckman, 1979](#); [Mroz, 1987](#)).

age unemployment rate at länder level.¹¹ Finally, we included missing indicators for each of the aforementioned variables in order to save observations and not to lose statistical power. Tables containing summary statistics of the control variables are in Appendix C.

We did not control for contemporaneous characteristics because their realization may be the consequence of the treatment and be therefore themselves outcome variables, i.e. “bad controls” (Angrist and Pischke, 2009, p. 64). For example, the educational level that an individual achieves may be causally impacted by past parental unemployment. Education is linked with earnings and income which, in turn, are used to buy health goods and improve the health condition (Grossman, 1972, 2000; Mincer, 1974).

4 Results

4.1 Main findings

The main findings are reported in Table 5. We estimated three models, one for each developmental stage (Ermisch and Francesconi, 2001; Ermisch et al., 2004; Cunha and Heckman, 2007; Francesconi et al., 2010). Column (1) reports the results for the baseline specification. It includes the set of pre-treatment characteristics used in the EB matching plus the controls for the local unemployment rates and the sets of dummy variables for the year when the individual was interviewed and born. Local unemployment rates are meant to control for the precariousness in the labor market at the period of when the treatment occurred. The survey year and the year of birth are meant to control for potential unobserved heterogeneity at time level, at both the childhood and the adulthood periods. Column (2) extends the baseline specification by adding other predetermined family characteristics, measured when the treatment occurred (i.e. individuals’ childhood), in order to reduce the risk of possible omitted variables bias: household size, house size in squared meters and the duration of parental unemployment. Finally, column (3) includes the lagged parental job loss expectations. Being the richest specification in terms of predetermined covariates, column (3) is our preferred specification.

We found mixed effects of parental unemployment on children’s health later in life. On the one hand, we detected a small and statistically insignificant effect on mental health.

¹¹Data come from German Employment Agency statistics. We took the average for the whole treatment period. We used national unemployment rates whenever data at länder level were not available. National level data come from the Federal Reserve of St. Louis economic data. National unemployment rates prior 1991 refer to West Germany.

On the other hand, we found a strong negative effect on physical health. The effect is stronger if parental unemployment occurred earlier (0-5 years old) or later (11-15 years old). Instead, if parental unemployment happened in the middle period (i.e. child between 6 and 10 years of age), the effect is small and not significantly different from 0. This non-linear pattern is similar to the one presented by [Nikolova and Nikolaev \(2021\)](#) for life satisfaction. These findings hold regardless of the model specification, even after controlling for parental job loss expectations.

Table 5: Main estimation results

	Outcome	Treated sample mean	Sample size	Baseline specification		With further family characteristics, not including parental job loss expectations		With further family characteristics, including also parental job loss expectations	
				(1)		(2)		(3)	
				ATT	R ²	ATT	R ²	ATT	R ²
Parental unemployment at 0-5	MCS	48.804	572	-2.268 (1.887)	0.495	-0.830 (2.158)	0.513	-0.818 (2.283)	0.529
	PCS	53.237	572	-4.506*** (1.333)	0.596	-4.612*** (1.535)	0.601	-4.658*** (1.603)	0.605
Parental unemployment at 6-10	MCS	49.100	1,335	0.887 (0.878)	0.265	-0.442 (1.011)	0.273	-0.882 (1.033)	0.289
	PCS	55.161	1,335	0.705 (0.620)	0.222	0.395 (0.715)	0.227	0.871 (0.733)	0.251
Parental unemployment at 11-15	MCS	49.855	1,693	0.739 (0.720)	0.349	1.401* (0.812)	0.353	1.381 (0.857)	0.357
	PCS	53.645	1,693	-1.966*** (0.512)	0.372	-2.252*** (0.575)	0.378	-2.597*** (0.603)	0.405

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Clustered standard errors at individual level are reported in parentheses. The baseline specification (column (1)) includes the following controls: individual age, gender, individual migration background, whether the individual is the firstborn child, average number of siblings, länder in which the individual was born, year of birth, survey year, paternal age when the child was born, paternal migration background, paternal educational level achieved, paternal cumulative unemployment, maternal age when the child was born, maternal migration background, maternal educational level achieved, maternal cumulative unemployment, average household income quintile and local unemployment rates during the treatment period. Column (2) adds to the previous specification average household size, average house size in squared meters and the duration of the unemployment spell experienced by the parent during individual's childhood. Finally, in column (3) we added also the lagged parental job loss expectations.

The point estimate of the effect of parental unemployment is -4.658 when it occurred when the child was between 0 and 5 years of age and -2.597 if it occurred in the latest period (i.e. at 11-15). The fact that the former effect of parental unemployment is in size almost twice as large as the latter suggests that early stages matter the most in the human capital production function ([Cunha and Heckman, 2007](#)). As early investments lead to higher returns later in life, early setbacks lead to greater losses. Moreover, the estimated effects on physical health are sizeable if compared to the treated average outcome: the occurrence of parental unemployment reduces physical health by 4.6%-8%. Among the regressors we also included the duration of the unemployment spell of the displaced parent. The effect on physical health of parental unemployment of average duration amounts to -5.971 (-2.639) if the displacement occurred when the child was between 0 and 5 (11

and 15) years of age.¹² Another way of quantifying the importance of the health effects of parental unemployment is to compare it to the health effect of aging: parental unemployment exerts an effect that is similar to the health effect induced by a 10-year aging of the individual. This suggests that it was the occurrence of unemployment that mattered most, rather than its duration. Hence, our results point towards a large detrimental role played by parental unemployment due to plant closure in shaping the individual physical health, with a scarring effect that lasts in the long-run.

Our findings may be the result of different mechanisms which took place during individuals' lives. The nil effect on mental health may be the consequence of a 'normalization' process that the laid-off parent operated over his/her unemployment condition (Georgellis et al., 2008; Thill et al., 2019). Individuals often use coping strategies to overcome stressful situations (Ashforth and Kreiner, 2002). A laid-off parent might have justified his/her condition by pointing to the impossibility of preventing an event that was out of his/her control, like a plant closure (Pignault and Houssemand, 2017, 2018). Moreover, when a generalized precarious conditions occur, individuals suffer less from their own misery because it is becoming the prevailing social norm (Clark, 2003, 2006; Clark et al., 2009; Stavrova et al., 2011). Normalizing his/her own unemployment might prevent displaced parents to develop anxiety or depression disorders, which are likely to affect offspring's well-being. The negative effect on physical health may be instead explained by a longer exposure of the child to parental risky behaviors. Laid-off workers are more likely to engage in unhealthy behaviors, like smoking or drinking (Schunck and Rogge, 2010; Reine et al., 2013; Mörk et al., 2020). Children with smoking or alcohol consumer parents are at higher risk of developing health problems while they are growing (Richter and Richter, 2001; Kuppens et al., 2020). If displaced workers are more likely to develop unhealthy habits, than a far-reaching effect of parental unemployment may be the capability of triggering unhealthy behaviors on children, impairing offspring's physical health.

We empirically tested whether parental unemployment may have consequences on children's physical health later in life by increasing their probability of engaging in risky behaviors. More in detail, we looked at the impact of parental unemployment on children's alcohol consumption, tobacco consumption and obesity when adult. We coded alcohol consumption as a dummy indicator equal to 1 if the individual had a moder-

¹²The full set of estimations results are reported in Appendix D.

ate/severe level of alcohol intakes.¹³ Tobacco consumption was defined by using a binary indicator which took value 1 if the individual reported to be a smoker at the time of the interview and 0 otherwise. Finally, obesity was defined by using a binary indicator which took value 1 if the individual reported a BMI score equal to or greater than 30, and 0 otherwise. We estimated three linear probability random-effect models for these three binary variables. Table 6 reports the results.

We found that parental unemployment increases the probability of engaging in risky behaviors later in life. Treated individuals are more likely to consume alcohol and tobacco. There are no significant differences in terms of obesity. Consistently with the findings in Table 5, the largest effect is for individuals who suffered parental unemployment in their early childhood (i.e. between 0 and 5 years of age). Individuals who experienced parent unemployment due to a plant closure when they were 0 to 5 years old are 46.0 and 27.3 percentage points more likely to frequently consume alcohol and to be smokers compared to their control peers. These findings are in line with those in Ermisch et al. (2004), who found that parental joblessness increases the likelihood of both smoking and experiencing greater distress later in life. Furthermore, these findings keep corroborating the idea that early stages represent a critical period in the child development (Cunha and Heckman, 2007).

Table 6: Mediating channels estimation results

	Moderate/severe alcohol consumption				Being a smoker				Obesity (BMI > 30)			
	(1)				(2)				(3)			
	<i>n</i>	ATT	σ	R^2	<i>n</i>	ATT	σ	R^2	<i>n</i>	ATT	σ	R^2
Parental unemployment at 0-5	236	0.460***	0.148	0.830	580	0.273**	0.127	0.756	571	0.051	0.074	0.379
Parental unemployment at 6-10	527	0.088	0.067	0.488	1,362	-0.087	0.055	0.451	1,344	-0.053*	0.030	0.226
Parental unemployment at 11-15	643	0.125**	0.061	0.476	1,725	0.068*	0.040	0.525	1,700	-0.019	0.026	0.376

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Standard errors (σ) are clustered at individual level. Model specification for each outcome variable and each treatment interval is the same as in column (3) of Table 5. Number of treated observations at 0-5 ages is 18 in column (1) and 35 in columns (2) and (3). Number of treated observations at 6-10 ages is 49 in column (1), 85 in column (2) and 84 in column (3). Number of treated observations at 11-15 ages is 36 in column (1) and 69 in columns (2) and (3).

¹³SOEP has four questions, each of which refers to a different type of drink (i.e. beer, wine, spirits and cocktails). The questions ask how often the interviewee consumes that specific alcoholic beverage. Four options are available: never, rarely, once in a while and often. We assigned value 1 to ‘never’ and 4 to ‘often’. We defined an individual to have a moderate/severe level of alcohol intake if the sum of the four scales was equal to or greater than 10.

4.2 Heterogeneity analysis

Parental unemployment may exert different effects according to the gender of either the parent or the child. We followed the previous empirical literature and tested this hypothesis in our sample (Haiken-DeNew and Kind, 2012; Powdthavee and Vernoit, 2013; Pieters and Rawlings, 2020; Briody, 2021; Nikolova and Nikolaev, 2021). First, we decomposed the treatment distinguishing between maternal and paternal unemployment. Second, we split each subsample according to the gender of the child. Tables 7 and 8 report the findings.

Table 7: Maternal and paternal unemployment effects at comparison

	MCS			PCS		
	(1)			(2)		
	<i>n</i>	ATT	<i>R</i> ²	<i>n</i>	ATT	<i>R</i> ²
Maternal unemployment at 0-5	572	-8.289** (3.298)	0.547	572	-10.052*** (2.336)	0.629
Paternal unemployment at 0-5		6.679** (3.305)			0.707 (2.330)	
Maternal unemployment at 6-10	1,335	-1.480 (1.134)	0.289	1,335	0.222 (0.804)	0.254
Paternal unemployment at 6-10		0.798 (1.667)			2.676** (1.178)	
Maternal unemployment at 11-15	1,693	3.060*** (1.047)	0.361	1,693	-1.136 (0.740)	0.405
Paternal unemployment at 11-15		-2.444** (1.200)			-3.674*** (0.850)	

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Standard errors are clustered at individual level and in parentheses. We split the plant closure treatment in two: maternal and paternal. Model specification includes both jointly. Model specification is the same as in column (3) of Table 5. Number of maternal unemployment observations is 17 at 0-5 ages, 70 at 6-10 ages and 34 at 11-15 aged in the full sample. Number of paternal unemployed observations is 16 at 0-5 ages, 14 at 6-10 ages and 36 at 11-15 ages in the full sample.

By looking at the estimates for mental health in column (1) of Table 7, we realised that the previous nil finding on mental health is the result of internal heterogeneity. When children are between 0 and 5 years of age, maternal (paternal) unemployment negatively (positively) affects children's mental health. The relation reverses if maternal and paternal unemployment occurred during the early adolescence (11-15 years of age). About the impact on physical health, we found that only the maternal unemployment matters during the early childhood, whereas only paternal unemployment is relevant during the early adolescence.

Table 8: Parental unemployment effects on daughters and sons

	MCS						PCS					
	Daughters			Sons			Daughters			Sons		
	(1a)			(2a)			(1b)			(2b)		
	<i>n</i>	ATT	<i>R</i> ²	<i>n</i>	ATT	<i>R</i> ²	<i>n</i>	ATT	<i>R</i> ²	<i>n</i>	ATT	<i>R</i> ²
Parental unemployment at 0-5	291	3.422 (10.754)	0.548	281	2.568 (13.944)	0.395	291	-12.722** (6.061)	0.715	281	-4.853 (14.033)	0.517
Parental unemployment at 6-10	709	2.917 (1.781)	0.388	626	-3.960* (2.142)	0.433	709	2.416* (1.255)	0.388	626	0.691 (1.625)	0.280
Parental unemployment at 11-15	823	-1.205 (1.616)	0.422	870	-1.288 (1.478)	0.498	823	-3.843*** (1.116)	0.535	870	0.268 (1.077)	0.345

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Standard errors are clustered at individual level and in parentheses. We split the plant closure treatment in two: maternal and paternal. Model specification includes both jointly. Model specification is the same as in column (3) of Table 5.

By splitting the sample according to the gender of child, we found that daughters suffer from parental unemployment in terms of physical health, whereas sons remain generally unaffected. Consistently with previous results, the largest effect comes from girls who were between 0 and 5 years old when experienced paternal unemployment. This piece of evidence is in contrast with the previous empirical literature, which found instead that are rather the sons who suffered the most from the negative labor market dynamics of the parents (Haiken-DeNew and Kind, 2012; Nikolova and Nikolaev, 2021). However, this literature focused on life satisfaction, whereas this paper is about different health outcomes.

4.3 Sensitivity checks with alternative matching procedures

In Section 3 we described the identification strategy. We combined matching methods and standard parametric estimation while exploiting an exogenous variation in the individual labor market condition to achieve the causal interpretation of the estimates. The matching was intended to increase the balance properties of the sample and reduce the potential model dependence of the subsequent estimates. In this section, we check the sensitivity of our results if we used more traditional balancing/matching algorithms based on the propensity score: the inverse probability weighting (IPW) and the kernel matching (KM).

The inverse probability weighting assigns to each control unit a weight equals to $\frac{Pr(\mathbf{X}_{it})}{1-Pr(\mathbf{X}_{it})}$, where $Pr(\mathbf{X}_{it})$ is the predicted probability of being treated. Treated units are instead assigned with weights equal to 1. The kernel matching creates counterfactuals for each treated observation by taking the average of those j control units that are similar in

terms of their propensity scores (Stuart, 2010).¹⁴ Matching techniques based on predicted probabilities suffer from the problem of ‘large propensity score values’, which may bias the subsequent results. We tried to limit this concern and, before estimating Equation (1), we retained only those observations within the common support (Peter, 2016). Table 9 reports the results for the sensitivity analysis.

Table 9: Sensitivity analysis for alternative matching methods

	Inverse probability weighting						Kernel matching					
	MCS			PCS			MCS			PCS		
	(1a)			(1b)			(2a)			(2b)		
	<i>n</i>	ATT	<i>R</i> ²	<i>n</i>	ATT	<i>R</i> ²	<i>n</i>	ATT	<i>R</i> ²	<i>n</i>	ATT	<i>R</i> ²
Parental unemployment at 0-5	886	-1.139 (2.354)	0.212	886	-3.233* (1.658)	0.209	1,263	-2.309 (1.626)	0.201	1,263	-3.365*** (1.132)	0.233
Parental unemployment at 6-10	2,449	1.155 (1.333)	0.115	2,449	1.082 (0.868)	0.144	2,982	1.813* (0.965)	0.118	2,982	1.107* (0.620)	0.130
Parental unemployment at 11-15	2,927	-1.005 (1.377)	0.089	2,927	-2.103** (0.904)	0.116	3,783	-1.121 (0.947)	0.107	3,783	-0.496 (0.630)	0.102

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Standard errors are clustered at individual level and in parentheses. Model specification in both matching techniques and for both outcome variables is the same as in column (3) of Table 5.

Replacing the CEM-EB algorithm with other balancing/matching techniques did not lead to relevant changes in the results. In fact, the findings on the non-linear pattern and the size of the estimates remained similar to those presented in Table 5. As before, there is a nil effect on mental health in all the subsamples. About physical health, there are sizable effects only for the 0-5 and 11-15 subsamples, with the former showing the largest impact. However, now the statistical significance of the estimates slightly fades away. This variation may be due to the less binding assumptions of both the IPW and KM matching. In fact, differently from the CEM-EB algorithm, neither the IPW nor the KM perform an exact matching nor requires identical moments conditions for the treated and control units.

5 Conclusions

In this paper, we estimated the impact of parental unemployment on children’s mental and physical health later in life. We assessed the causal effect of parental unemployment by using plant closures as exogenous variation of the individual labor market condition

¹⁴We used the Stata command `psmatch2`. Kernel matching is performed by using the Epanechnikov function.

([Brand, 2015](#)). We combined it with bias-adjusted methods robust to model misspecification to strengthen the causal interpretation of the estimates ([Ho et al., 2007](#)). Information on parental unemployment referred to the period when individuals were children (i.e. 0 to 15 years old). We split the childhood period in three developmental stages in order to test whether parental unemployment exerts different effects depending on the timing of its occurrence ([Ermisch and Francesconi, 2001](#); [Ermisch et al., 2004](#); [Cunha and Heckman, 2007](#)). Information on health outcomes referred to the period when individuals were adults (i.e. 18 to 31 years old). We used data from the German Socio-Economic Panel from 2002 until 2018.

We found that parental unemployment exerts mixed effects on individuals' health later in life. On the one hand, mental health was unaffected. This finding holds regardless of the model specification used and the timing of the occurrence of the treatment. On the other hand, we found that parental unemployment exerts a negative and significant effect on physical health. The largest impact is generated when parental unemployment occurred during the early childhood (i.e. between 0 and 5 years of age). This effect is almost twice in size compared to the effect found for the case in which parental unemployment occurred during the early adolescence (i.e. between 11 and 15 years of age). This is consistent with theoretical predictions which indicate the early periods of life are the most important in terms of human capital accumulation ([Cunha and Heckman, 2007](#)). We found that the effect is heterogeneous across gender, both in terms of the child experiencing parental unemployment and in terms of the gender of the displaced parent. Maternal unemployment mattered initially, whereas paternal unemployment mattered later. Daughters suffered the negative consequences of parental unemployment, whilst sons remained generally unaffected. We also found that parental unemployment determines a higher probability of smoking and drinking later in life, which may be some of the mediating channels through which the negative impact on physical health was triggered.

Our findings raise important policy considerations. First, we provided evidence of long-term scarring effects of parental unemployment on children's health. Most of the previous empirical research discussed only short-term effects, while neglecting possible consequences in the long-run. Policy-makers should be aware that parental unemployment generates long-lasting negative effects on physical health and that intervention is needed to alleviate them. Second, we found that most of the effect is due to the occurrence of the treatment, rather than its duration. This result is in line with previous empirical research and meta-analyses on the health effects of unemployment ([Cygan-Rehm](#)

et al., 2017; Picchio and Ubaldi, 2022). Policy-makers may consider the need for therapeutic support for households that experienced an episode of parental unemployment, promptly after the job loss event. Third, we found that parental unemployment generates higher chances of smoking and drinking later in life. Hence, policy-makers may also want to consider tax increases on alcohol beverages and tobacco in order to discourage people, especially the unemployed, to initiate the consumption since the very beginning (Chaloupka et al., 2002; Elder et al., 2010; Chaloupka et al., 2011).

Finally, our results have some limitations and should be qualified for some reasons. First, we studied an intergenerational phenomenon by relying on survey data and self-reported measures of health. In this kind of analysis, administrative records should be preferred, because they are less likely to be plagued by measurement errors and may provide more objective measures of health. Second, even though we exploited an exogenous variation in the individual labor market condition and used it within a double-robust framework to identify the causal effect of the parental unemployment, it could still be possible that there were other confounding unobserved components which may undermine the reliability of our results. Third, studying the long-run consequences of parental unemployment experienced during childhood on health outcomes during adulthood is demanding in terms of number of observations for a survey data, because an individual should be observed both during her/his childhood and during her/his adulthood. This resulted in a large loss of observations, in a reduction in the representativeness of the final sample and in a loss of statistical power of the empirical analysis.

References

- Andersen, H. H., Mühlbacher, A., Nübling, M., Schupp, J. and Wagner, G. G. (2007). Computation of standard values for physical and mental health scale scores using the SOEP version of the SF-12v2, *Journal of Contextual Economics – Schmollers Jahrbuch* **127**(1): 171–182.
- Angrist, J. and Pischke, J. (2009). *Mostly harmless econometrics: an empiricist's companion*, Princeton University Press, Princeton.
- Arulampalam, W. (2001). Is unemployment really scarring? Effects of unemployment experiences on wages, *The Economic Journal* **111**(475): F585–F606.
- Ashforth, B. E. and Kreiner, G. E. (2002). Normalizing emotion in organizations: Making the extraordinary seem ordinary, *Human Resource Management Review* **12**(2): 215–235.

- Bantle, C. and Haiken-DeNew, J. P. (2002). Smoke signals: The intergenerational transmission of smoking behavior. DIW Discussion Paper No. 277, DIW Berlin.
- Becker, G. S. and Tomes, N. (1979). An equilibrium theory of the distribution of income and intergenerational mobility, *Journal of Political Economy* **87**(6): 1153–1189.
- Becker, G. S. and Tomes, N. (1986). Human capital and the rise and fall of families, *Journal of Labor Economics* **4**(3, Part 2): S1–S39.
- Blackwell, M., Iacus, S., King, G. and Porro, G. (2009). cem: Coarsened exact matching in Stata, *The Stata Journal* **9**(4): 524–546.
- Brand, J. E. (2015). The far-reaching impact of job loss and unemployment, *Annual Review of Sociology* **41**: 359.
- Brand, J. E. and Thomas, J. S. (2014). Job displacement among single mothers: Effects on children's outcomes in young adulthood, *American Journal of Sociology* **119**(4): 955–1001.
- Bratberg, E., Nilsen, Ø. A. and Vaage, K. (2008). Job losses and child outcomes, *Labour Economics* **15**(4): 591–603.
- Briody, J. (2021). Parental unemployment during the Great Recession and childhood adiposity, *Social Science & Medicine* **275**: 113798.
- Chaloupka, F. J., Grossman, M. and Saffer, H. (2002). The effects of price on alcohol consumption and alcohol-related problems, *Alcohol Research & Health* **26**(1): 22–34.
- Chaloupka, F. J., Straif, K. and Leon, M. E. (2011). Effectiveness of tax and price policies in tobacco control, *Tobacco Control* **20**(3): 235–238.
- Chan, S. and Stevens, A. H. (2001). Job loss and employment patterns of older workers, *Journal of Labor Economics* **19**(2): 484–521.
- Clark, A. (2006). A note on unhappiness and unemployment duration. IZA Discussion Paper No. 2406, IZA Institute of Labor Economics, Bonn.
- Clark, A. E. (2003). Unemployment as a social norm: Psychological evidence from panel data, *Journal of Labor Economics* **21**(2): 323–351.
- Clark, A., Knabe, A. and Rätzl, S. (2009). Unemployment as a social norm in Germany, *Journal of Contextual Economics – Schmollers Jahrbuch* **129**(2): 251–260.
- Coelli, M. (2011). Parental job loss and the education enrollment of youth, *Labour Economics* **18**(1): 25–35.
- Cunha, F. and Heckman, J. (2007). The technology of skill formation, *American Economic Review* **97**(2): 31–47.

- Cygan-Rehm, K., Kuehnle, D. and Oberfichtner, M. (2017). Bounding the causal effect of unemployment on mental health: Nonparametric evidence from four countries, *Health Economics* **26**(12): 1844–1861.
- Di Nallo, A., Lipps, O., Oesch, D. and Voorpostel, M. (2022). The effect of unemployment on couples separating in Germany and the UK, *Journal of Marriage and Family* **84**(1): 310–329.
- Doepke, M., Sorrenti, G. and Zilibotti, F. (2019). The economics of parenting, *Annual Review of Economics* **11**: 55–84.
- Drydakis, N. (2015). The effect of unemployment on self-reported health and mental health in Greece from 2008 to 2013: A longitudinal study before and during the financial crisis, *Social Science & Medicine* **128**: 43–51.
- Ekhaugen, T. (2009). Extracting the causal component from the intergenerational correlation in unemployment, *Journal of Population Economics* **22**(1): 97–113.
- Elder, R. W., Lawrence, B., Ferguson, A., Naimi, T. S., Brewer, R. D., Chattopadhyay, S. K., Toomey, T. L., Fielding, J. E., on Community Preventive Services, T. F. et al. (2010). The effectiveness of tax policy interventions for reducing excessive alcohol consumption and related harms, *American Journal of Preventive Medicine* **38**(2): 217–229.
- Eliason, M. (2012). Lost jobs, broken marriages, *Journal of Population Economics* **25**(4): 1365–1397.
- Ermisch, J. F. and Francesconi, M. (2001). Family structure and children’s achievements, *Journal of Population Economics* **14**(2): 249–270.
- Ermisch, J. and Francesconi, M. (2013). The effect of parental employment on child schooling, *Journal of Applied Econometrics* **28**(5): 796–822.
- Ermisch, J., Francesconi, M. and Pevalin, D. J. (2004). Parental partnership and joblessness in childhood and their influence on young people’s outcomes, *Journal of the Royal Statistical Society: Series A (Statistics in Society)* **167**(1): 69–101.
- Farré, L., Fasani, F. and Mueller, H. (2018). Feeling useless: the effect of unemployment on mental health in the Great Recession, *IZA Journal of Labor Economics* **7**(1): 1–34.
- Filomena, M. (2023). Unemployment scarring effects: An overview and meta-analysis of empirical studies, *Italian Economic Journal* (forthcoming).
- Francesconi, M., Jenkins, S. P. and Siedler, T. (2010). Childhood family structure and schooling outcomes: Evidence for Germany, *Journal of Population Economics* **23**(3): 1073–1103.
- Georgellis, Y., Gregoriou, A., Healy, J. and Tsitsianis, N. (2008). Unemployment and life satisfaction: A non-linear adaptation process, *International Journal of Manpower* **29**(7): 668–680.

- Göhlmann, S., Schmidt, C. M. and Tauchmann, H. (2010). Smoking initiation in germany: the role of intergenerational transmission, *Health Economics* **19**(2): 227–242.
- Gill, S. C., Butterworth, P., Rodgers, B. and Mackinnon, A. (2007). Validity of the mental health component scale of the 12-item Short-Form Health Survey (MCS-12) as measure of common mental disorders in the general population, *Psychiatric Research* **152**(1): 63–71.
- Goebel, J., Grabka, M. M., Liebig, S., Kroh, M., Richter, D., Schröder, C. and Schupp, J. (2019). The German Socio-Economic Panel (soep), *Jahrbücher für Nationalökonomie und Statistik* **239**(2): 345–360.
- Green, F. (2011). Unpacking the misery multiplier: How employability modifies the impacts of unemployment and job insecurity on life satisfaction and mental health, *Journal of Health Economics* **30**(2): 265–276.
- Grossman, M. (1972). On the concept of health capital and the demand for health, *Journal of Political Economy* **80**(2): 223–266.
- Grossman, M. (2000). The human capital model, in A. Culyer and J. Newhouse (eds), *Handbook of Health Economics*, Vol. 1, Part A, North-Holland, Amsterdam, pp. 347–408.
- Hainmueller, J. (2012). Entropy balancing for causal effects: A multivariate reweighting method to produce balanced samples in observational studies, *Political Analysis* **20**(1): 25–46.
- Hainmueller, J. and Xu, Y. (2013). Ebalance: A Stata package for entropy balancing, *Journal of Statistical Software* **54**(7).
- Haisken-DeNew, J. P. and Kind, M. (2012). Unexpected victims: How parents’ unemployment affects their children’s life satisfaction. Melbourne Institute Working Paper No. 2/12, Melbourne Institute of Applied Economic and Social Research.
- Headey, B., Kelley, J. and Wearing, A. (1993). Dimensions of mental health: Life satisfaction, positive affect, anxiety and depression, *Social Indicators Research* **29**(1): 63–82.
- Heckman, J. J. (1979). Sample selection bias as a specification error, *Econometrica* **47**(1): 153–161.
- Hérault, N. and Kalb, G. (2016). Intergenerational correlation of labor market outcomes, *Review of Economics of the Household* **14**(1): 231–249.
- Ho, D. E., Imai, K., King, G. and Stuart, E. A. (2007). Matching as nonparametric preprocessing for reducing model dependence in parametric causal inference, *Political Analysis* **15**(3): 199–236.
- Iacus, S. M., King, G. and Porro, G. (2008). Matching for causal inference without balance checking, *Available at SSRN 1152391*.
- Iacus, S. M., King, G. and Porro, G. (2011). Multivariate matching methods that are monotonic imbalance bounding, *Journal of the American Statistical Association* **106**(493): 345–361.

- Jacobson, L. S., LaLonde, R. J. and Sullivan, D. G. (1993). Earnings losses of displaced workers, *The American Economic Review* pp. 685–709.
- Jahoda, M. (1982). *Employment and unemployment: A social-psychological analysis*, New York: Cambridge University Press.
- Janlert, U. and Hammarström, A. (2009). Which theory is best? Explanatory models of the relationship between unemployment and health, *BMC Public Health* **9**(1): 1–9.
- Jensen, P. and Smith, N. (1990). Unemployment and marital dissolution, *Journal of Population Economics* **3**(3): 215–229.
- Jones, A. M., Rice, N. and Zantomio, F. (2020). Acute health shocks and labour market outcomes: Evidence from the post crash era, *Economics & Human Biology* **36**: 100811.
- Kalil, A. and Ziol-Guest, K. M. (2008). Parental employment circumstances and children’s academic progress, *Social Science Research* **37**(2): 500–515.
- Karhula, A., Lehti, H. and Jani, E. (2017). Intergenerational scars? The long-term effects of parental unemployment during a depression, *Research on Finnish Society* **10**: 87–99.
- Kassenboehmer, S. C. and Haisken-DeNew, J. P. (2009). You’re fired! The causal negative effect of entry unemployment on life satisfaction, *The Economic Journal* **119**(536): 448–462.
- Knabe, A., Rätzl, S., Schöb, R. and Weimann, J. (2010). Dissatisfied with life but having a good day: Time-use and well-being of the unemployed, *The Economic Journal* **120**(547): 867–889.
- Kuppens, S., Moore, S. C., Gross, V., Lowthian, E. and Siddaway, A. P. (2020). The enduring effects of parental alcohol, tobacco, and drug use on child well-being: A multilevel meta-analysis, *Development and Psychopathology* **32**(2): 765–778.
- Lam, J. and Ambrey, C. L. (2019). The scarring effects of father’s unemployment? Job-security satisfaction and mental health at midlife, *The Journals of Gerontology: Series B* **74**(1): 105–112.
- Liang, K. and Zeger, S. (1986). Longitudinal data analysis using generalized linear models, *Biometrika* **73**(1): 13–22.
- Lindo, J. M. (2011). Parental job loss and infant health, *Journal of Health Economics* **30**(5): 869–879.
- Macmillan, L. (2014). Intergenerational worklessness in the UK and the role of local labor markets, *Oxford Economic Papers* **66**(3): 871–889.
- Marcus, J. (2013). The effect of unemployment on the mental health of spouses – Evidence from plant closures in Germany, *Journal of Health Economics* **32**(3): 546–558.

- Masten, A. S. and Narayan, A. J. (2012). Child development in the context of disaster, war, and terrorism: Pathways of risk and resilience, *Annual Review of Psychology* **63**: 227.
- Masten, A. S. and Palmer, A. R. (2019). Parenting to promote resilience in children, *Handbook of Parenting*, Routledge, pp. 156–188.
- Mincer, J. (1974). *Schooling, experience, and earnings*, National Bureau of Economic Research.
- Mörk, E., Sjögren, A. and Svaleryd, H. (2020). Consequences of parental job loss on the family environment and on human capital formation – Evidence from workplace closures, *Labour Economics* **67**: 101911.
- Mroz, T. A. (1987). The sensitivity of an empirical model of married women's hours of work to economic and statistical assumptions, *Econometrica* **55**(4): 765–799.
- Müller, S., Riphahn, R. T. and Schwientek, C. (2017). Paternal unemployment during childhood: Causal effects on youth worklessness and educational attainment, *Oxford Economic Papers* **69**(1): 213–238.
- Neubert, M., Süßenbach, P., Rief, W. and Euteneuer, F. (2019). Unemployment and mental health in the German population: The role of subjective social status, *Psychology Research and Behavior Management* **12**: 557–564.
- Nikolova, M. and Ayhan, S. H. (2019). Your spouse is fired! How much do you care?, *Journal of Population Economics* **32**(3): 799–844.
- Nikolova, M. and Nikolaev, B. N. (2021). Family matters: The effects of parental unemployment in early childhood and adolescence on subjective well-being later in life, *Journal of Economic Behavior & Organization* **181**: 312–331.
- Nübling, M., Andersen, H. H. and Mühlbacher, A. (2006). Entwicklung eines verfahrens zur berechnung der körperlichen und psychischen summenskalen auf basis der SOEP – version des SF 12 (algorithmus), *Technical report*, DIW Data Documentation.
- Oreopoulos, P., Page, M. and Stevens, A. H. (2008). The intergenerational effects of worker displacement, *Journal of Labor Economics* **26**(3): 455–483.
- Peter, F. (2016). The effect of involuntary maternal job loss on children's behaviour and non-cognitive skills, *Labour Economics* **42**: 43–63.
- Picchio, M. and Ubaldi, M. (2022). Unemployment and health: A meta-analysis. IZA Discussion Paper No. 15433, IZA Institute of Labor Economics, Bonn.
- Pieters, J. and Rawlings, S. (2020). Parental unemployment and child health in China, *Review of Economics of the Household* **18**(1): 207–237.
- Pignault, A. and Houssemand, C. (2017). Normalizing unemployment: A new way to cope with unemployment?, *Basic and Applied Social Psychology* **39**(6): 372–386.

- Pignault, A. and Houssemand, C. (2018). An alternative relationship to unemployment: Conceptualizing unemployment normalization, *Review of General Psychology* **22**(3): 355–366.
- Powdthavee, N. and Vernoit, J. (2013). Parental unemployment and children's happiness: A longitudinal study of young people's well-being in unemployed households, *Labour Economics* **24**: 253–263.
- Rege, M., Telle, K. and Votruba, M. (2011). Parental job loss and children's school performance, *Review of Economic Studies* **78**(4): 1462–1489.
- Reine, I., Novo, M. and Hammarström, A. (2013). Unemployment and ill health – A gender analysis: Results from a 14-year follow-up of the northern Swedish cohort, *Public Health* **127**(3): 214–222.
- Richter, L. and Richter, D. M. (2001). Exposure to parental tobacco and alcohol use: Effects on children's health and development, *American Journal of Orthopsychiatry* **71**(2): 182–203.
- Schmidt, C. M. and Tauchmann, H. (2011). Heterogeneity in the intergenerational transmission of alcohol consumption: A quantile regression approach, *Journal of Health Economics* **30**(1): 33–42.
- Schmitz, H. (2011). Why are the unemployed in worse health? The causal effect of unemployment on health, *Labour Economics* **18**(1): 71–78.
- Schunck, R. and Rogge, B. G. (2010). Unemployment and its association with health-relevant actions: Investigating the role of time perspective with German census data, *International Journal of Public Health* **55**(4): 271–278.
- Siedler, T. (2011). Parental unemployment and young people's extreme party affinity: Evidence from panel data, *Journal of Royal Statistical Society: Series A (Statistics in Society)* **174**(3): 737–758.
- Siegers, R., Steinhauer, H. W. and Dührsen, L. (2021). SOEP-Core v36: Documentation of sample sizes and panel attrition in the German Socio-Economic Panel (SOEP)(1984 until 2019). SOEP Survey Papers No. 960, DIW Berlin.
- Stauder, J. (2019). Unemployment, unemployment duration, and health: Selection or causation?, *European Journal of Health Economics* **20**(1): 59–73.
- Stavrova, O., Schlösser, T. and Fetchenhauer, D. (2011). Are the unemployed equally unhappy all around the world? The role of the social norms to work and welfare state provision in 28 OECD countries, *Journal of Economic Psychology* **32**(1): 159–171.
- Stevens, A. H. and Schaller, J. (2011). Short-run effects of parental job loss on children's academic achievement, *Economics of Education Review* **30**(2): 289–299.
- Stuart, E. A. (2010). Matching methods for causal inference: A review and a look forward, *Statistical Science* **25**(1): 1–21.

- Thill, S., Houssemand, C. and Pignault, A. (2019). Unemployment normalization: Its effect on mental health during various stages of unemployment, *Psychological Reports* **122**(5): 1600–1617.
- Wagner, G. G., Frick, J. R. and Schupp, J. (2007). The German Socio-Economic Panel study (SOEP)- evolution, scope and enhancements. SOEPpaper on Multidisciplinary Panel Data Research No. 1, DIW Berlin.
- Ware, J. E., Kosinski, M. and Dewey, J. (2000). How to score version 2 of the SF-36 Health Survey. Lincoln (RI): QualityMetric.
- Ware, J. E., Kosinski, M. and Keller, S. D. (1996). A 12-Item Short-Form Health Survey: Construction of scales and preliminary tests of reliability and validity, *Medical Care* **34**(3): 220–233.
- Winkelmann, L. and Winkelmann, R. (1998). Why are the unemployed so unhappy? Evidence from panel data, *Economica* **65**(257): 1–15.
- Yu, J. (2003). The association between parental alcohol-related behaviors and children's drinking, *Drug and Alcohol Dependence* **69**(3): 253–262.

Appendix

A Entropy balancing matching

Table A.1: Pre- and post-matching bias on the mean (treatment 0 to 5 years old)

Variables	Pre-Matching				Post-Matching			
	Mean Treated	Mean Controls	Difference	Bias	Mean Treated	Mean Controls	Difference	Bias
<i>Individual characteristics</i>								
Age	23.500	23.260	0.240	1.021%	23.500	23.439	0.061	0.260%
Sex (1 if female)	0.609	0.482	0.127	20.854%	0.609	0.600	0.009	1.478%
Migration background	0.054	0.015	0.039	72.223%	0.054	0.046	0.008	14.81%
Firstborn:								
Missing	0.103	0.138	-0.035	-33.980%	0.103	0.101	0.002	1.941%
Yes	0.277	0.392	-0.115	-41.516%	0.277	0.291	-0.014	-5.054%
Average number of siblings	1.092	0.836	0.256	23.443%	1.092	1.072	0.020	1.831%
Region where born:								
Missing	-	-	-	-	-	-	-	-
North Germany	0.000	0.094	-0.094	-	0.000	0.020	-0.020	-
South Germany	0.054	0.192	-0.138	-255.556%	0.054	0.089	-0.035	-64.815%
Central-West Germany	0.000	0.245	-0.245	-	0.000	0.045	-0.045	-
Central-East Germany	0.609	0.361	0.248	40.722%	0.609	0.548	0.061	10.016%
<i>Paternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.310	0.222	0.088	28.387%	0.310	0.278	0.032	10.323%
21 to 25 years old	0.342	0.351	-0.009	-2.632%	0.342	0.356	-0.014	-4.094%
26 to 30 years old	0.571	0.723	-0.152	-26.620%	0.571	0.587	-0.016	-2.802%
31 to 35 years old	0.375	0.265	0.110	29.334%	0.375	0.367	0.008	2.134%
36 to 40 years old	-	-	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-	-	-
Migration background	0.038	0.006	0.032	84.210%	0.038	0.031	0.007	18.421%
Educational level achieved:								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.815	0.891	-0.076	-9.325%	0.815	0.827	-0.012	-1.472%
High school diploma	0.076	0.020	0.056	73.684%	0.076	0.062	0.014	18.421%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.445	0.153	0.292	65.618%	0.445	0.403	0.042	9.438%
<i>Maternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.413	0.403	0.010	2.421%	0.413	0.401	0.012	2.905%
21 to 25 years old	0.370	0.534	-0.164	-44.324%	0.370	0.418	-0.048	-12.973%
26 to 30 years old	-	-	-	-	-	-	-	-
31 to 35 years old	-	-	-	-	-	-	-	-
36 to 40 years old	-	-	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-	-	-
Migration background	-	-	-	-	-	-	-	-
Educational level achieved:								
Missing	-	-	-	-	-	-	-	-
Less than HS	-	-	-	-	-	-	-	-
High school diploma	0.152	0.058	0.094	61.842%	0.152	0.128	0.024	15.789%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.252	0.371	-0.119	-47.223%	0.252	0.279	-0.027	-10.714%
<i>Household characteristics</i>								
HH income quintile:								
Missing	0.375	0.235	0.140	37.334%	0.375	0.349	0.026	6.934%
1st quintile	0.076	0.230	-0.154	-202.632%	0.076	0.098	-0.022	-28.947%
2nd quintile	0.277	0.219	0.053	19.485%	0.277	0.275	0.002	0.722%
3rd quintile	0.120	0.181	-0.061	-50.834%	0.120	0.136	-0.016	-13.334%
4th quintile	0.000	0.065	-0.065	-	0.000	0.010	-0.010	-
5th quintile	-	-	-	-	-	-	-	-

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: (-) means that the variable was omitted from the matching because of perfect collinearity. The various länders are grouped in four macro-regions. 'North Germany' includes Schleswig-Holstein, Hamburg, Mecklenburg-Vorpommern, Niedersachsen and Bremen. 'South Germany' includes Bayern and Baden-Wuerttemberg. 'Central-West Germany' includes Nordrhein-Westfalen, Hessen, Rheinland-Pfalz and Saarland. Finally, 'Central-East Germany' includes Berlin, Brandenburg, Sachsen, Sachsen-Anhalt and Thuringen.

Table A.2: Pre- and post-matching bias on the variance (treatment 0 to 5 years old)

Variables	Pre-Matching				Post-Matching			
	Variance Treated	Variance Controls	Difference	Bias %	Variance Treated	Variance Controls	Difference	Bias %
<i>Individual characteristics</i>								
Age	14.186	14.417	-0.231	-1.628%	14.186	14.188	-0.002	-0.014%
Sex (1 if female)	0.239	0.250	-0.011	-4.602%	0.239	0.240	-0.001	-0.418%
Migration background	0.052	0.014	0.038	73.077%	0.052	0.044	0.008	15.385%
Firstborn:								
Missing	0.093	0.119	-0.026	-27.957%	0.093	0.091	0.002	2.150%
Yes	0.201	0.238	-0.037	-18.408%	0.201	0.206	-0.005	-2.488%
Average number of siblings	0.587	0.494	0.093	15.843%	0.587	0.584	0.003	0.511%
Region where born:								
Missing	-	-	-	-	-	-	-	-
North Germany	0.000	0.085	-0.085	-	0.000	0.020	-0.020	-
South Germany	0.052	0.155	-0.103	-198.077%	0.052	0.081	-0.029	-55.769%
Central-West Germany	0.000	0.185	-0.185	-	0.000	0.043	-0.043	-
Central-East Germany	0.239	0.231	0.008	3.347%	0.239	0.248	-0.009	-3.766%
<i>Paternal characteristics</i>								
Age when child was born:								
Missing								
20 years old or less	0.215	0.173	0.042	19.535%	0.215	0.201	0.014	6.512%
21 to 25 years old	0.226	0.228	-0.002	-0.885%	0.226	0.229	-0.003	-1.327%
26 to 30 years old	0.246	0.200	0.046	18.699%	0.246	0.242	0.004	1.626%
31 to 35 years old	0.236	0.195	0.041	17.373%	0.236	0.232	0.004	1.695%
36 to 40 years old	-	-	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-	-	-
Migration background	0.037	0.005	0.032	86.486%	0.037	0.030	0.007	18.919%
Educational level achieved:								
Missing								
Less than HS	0.151	0.097	0.054	35.762%	0.151	0.143	0.008	5.298%
High school diploma	0.071	0.019	0.052	73.239%	0.071	0.058	0.013	18.310%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.532	0.163	0.369	69.361%	0.532	0.489	0.043	8.083%
<i>Maternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.244	0.241	0.003	1.229%	0.244	0.240	0.004	1.639%
21 to 25 years old	0.234	0.249	-0.015	-6.410%	0.234	0.243	-0.009	-3.846%
26 to 30 years old	-	-	-	-	-	-	-	-
31 to 35 years old	-	-	-	-	-	-	-	-
36 to 40 years old	-	-	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-	-	-
Migration background	-	-	-	-	-	-	-	-
Educational level achieved:								
Missing	-	-	-	-	-	-	-	-
Less than HS	-	-	-	-	-	-	-	-
High school diploma	0.130	0.054	0.076	58.461%	0.130	0.112	0.018	13.846%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.091	0.405	-0.314	-345.055%	0.091	0.149	-0.058	-63.736%
<i>Household characteristics</i>								
HH income quintile:								
Missing	0.236	0.180	0.056	23.729%	0.236	0.227	0.009	3.814%
1st quintile	0.071	0.177	-0.106	-149.296%	0.071	0.088	-0.017	-23.944%
2nd quintile	0.201	0.171	0.030	14.925%	0.201	0.200	0.001	0.497%
3rd quintile	0.106	0.148	-0.042	-39.623%	0.106	0.118	-0.012	-11.321%
4th quintile	0.000	0.061	-0.061	-	0.000	0.010	-0.010	-
5th quintile	-	-	-	-	-	-	-	-

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: (-) means that the variable was omitted from the matching because of perfect collinearity. The various länder are grouped in four macro-regions. 'North Germany' includes Schleswig-Holstein, Hamburg, Mecklenburg-Vorpommern, Niedersachsen and Bremen. 'South Germany' includes Bayern and Baden-Wuerttemberg. 'Central-West Germany' includes Nordrhein-Westfalen, Hessen, Rheinland-Pfalz and Saarland. Finally, 'Central-East Germany' includes Berlin, Brandenburg, Sachsen, Sachsen-Anhalt and Thuringen.

Table A.3: Pre- and post-matching bias on the mean (treatment 6 to 10 years old)

Variables	Pre-Matching				Post-Matching			
	Variance Treated	Variance Controls	Difference	Bias %	Variance Treated	Variance Controls	Difference	Bias %
<i>Individual characteristics</i>								
Age	24.000	23.346	0.654	2.725%	24.000	23.835	0.165	0.687%
Sex (1 if female)	0.720	0.508	0.212	29.445%	0.720	0.673	0.047	6.528%
Migration background	0.269	0.134	0.135	50.186%	0.269	0.251	0.018	6.691%
Firstborn:								
Missing	0.129	0.112	0.017	13.178%	0.129	0.124	0.005	3.876%
Yes	0.254	0.398	-0.144	-56.693%	0.254	0.286	-0.032	-12.598%
Average number of siblings	1.274	1.224	0.050	3.925%	1.274	1.264	0.010	0.785%
Region where born:								
Missing	-	-	-	-	-	-	-	-
North Germany	0.000	0.082	-0.082	-	0.000	0.026	-0.026	-
South Germany	0.095	0.160	-0.065	-68.421%	0.095	0.113	-0.018	-18.947%
Central-West Germany	0.112	0.195	-0.083	-74.107%	0.112	0.125	-0.013	-11.607%
Central-East Germany	0.750	0.501	0.249	33.200%	0.750	0.688	0.062	8.267%
<i>Paternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.050	0.016	0.034	68.000%	0.050	0.044	0.006	12.000%
21 to 25 years old	0.437	0.260	0.177	40.503%	0.437	0.407	0.030	6.865%
26 to 30 years old	0.312	0.498	-0.186	-59.615%	0.312	0.339	-0.027	-8.654%
31 to 35 years old	0.127	0.193	-0.066	-51.968%	0.127	0.144	-0.017	-13.386%
36 to 40 years old	0.056	0.031	0.025	44.643%	0.056	0.050	0.006	10.714%
41 years old or more	-	-	-	-	-	-	-	-
Migration background	0.263	0.131	0.132	50.190%	0.263	0.246	0.017	6.464%
Educational level achieved:								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.118	0.019	0.099	83.898%	0.118	0.104	0.014	11.864%
High school diploma	-	-	-	-	-	-	-	-
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.490	0.287	0.203	41.429%	0.490	0.487	0.003	0.612%
<i>Maternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.080	0.038	0.042	52.500%	0.080	0.072	0.008	10.000%
21 to 25 years old	0.558	0.410	0.148	26.523%	0.558	0.530	0.028	5.018%
26 to 30 years old	0.252	0.512	-0.260	-103.175%	0.252	0.300	-0.048	-19.048%
31 to 35 years old	0.050	0.027	0.023	46.000%	0.050	0.045	0.005	10.000%
36 to 40 years old	-	-	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-	-	-
Migration background	0.269	0.137	0.132	49.071%	0.269	0.252	0.017	6.320%
Educational level achieved:								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.134	0.064	0.070	52.239%	0.134	0.125	0.009	6.716%
High school diploma	0.746	0.825	-0.079	-10.590%	0.746	0.752	-0.006	-0.804%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.700	0.538	0.162	23.143%	0.700	0.664	0.036	5.143%
<i>Household characteristics</i>								
HH income quintile:								
Missing	-	-	-	-	-	-	-	-
1st quintile	0.243	0.288	-0.045	-18.518%	0.243	0.254	-0.011	-4.527%
2nd quintile	0.252	0.287	-0.035	-13.889%	0.252	0.263	-0.011	-4.365%
3rd quintile	0.392	0.235	0.157	40.051%	0.392	0.352	0.040	10.204%
4th quintile	0.019	0.045	-0.026	-136.842%	0.019	0.028	-0.009	-47.368%
5th quintile	-	-	-	-	-	-	-	-

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: (-) means that the variable was omitted from the matching because of perfect collinearity. The various länder are grouped in four macro-regions. 'North Germany' includes Schleswig-Holstein, Hamburg, Mecklenburg-Vorpommern, Niedersachsen and Bremen. 'South Germany' includes Bayern and Baden-Wuerttemberg. 'Central-West Germany' includes Nordrhein-Westfalen, Hessen, Rheinland-Pfalz and Saarland. Finally, 'Central-East Germany' includes Berlin, Brandenburg, Sachsen, Sachsen-Anhalt and Thuringen.

Table A.4: Pre- and post-matching bias on the variance (treatment 6 to 10 years old)

Variables	Pre-Matching				Post-Matching			
	Variance Treated	Variance Controls	Difference	Bias %	Variance Treated	Variance Controls	Difference	Bias %
<i>Individual characteristics</i>								
Age	15.637	14.346	1.291	8.256%	15.637	15.500	0.137	0.876%
Sex (1 if female)	0.202	0.250	-0.048	-23.762%	0.202	0.220	-0.018	-8.911%
Migration background	0.197	0.116	0.081	41.117%	0.197	0.188	0.009	4.568%
Firstborn:								
Missing	0.113	0.099	0.014	12.389%	0.113	0.108	0.005	4.425%
Yes	0.190	0.240	-0.050	-26.316%	0.190	0.204	-0.014	-7.368%
Average number of siblings	0.657	0.795	-0.138	-21.004%	0.657	0.701	-0.044	-6.697%
Region where born:								
Missing	-	-	-	-	-	-	-	-
North Germany	0.000	0.075	-0.075	-	0.000	0.026	-0.026	-
South Germany	0.086	0.134	-0.048	-55.814%	0.086	0.100	-0.014	-16.279%
Central-West Germany	0.100	0.157	-0.057	-57.000%	0.100	0.109	-0.009	-9.000%
Central-East Germany	0.188	0.250	-0.062	-32.979%	0.188	0.215	-0.027	-14.362%
<i>Paternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.047	0.016	0.031	65.957%	0.047	0.042	0.005	10.638%
21 to 25 years old	0.247	0.193	0.054	21.862%	0.247	0.241	0.006	2.429%
26 to 30 years old	0.215	0.250	-0.035	-16.279%	0.215	0.224	-0.009	-4.186%
31 to 35 years old	0.111	0.156	-0.045	-40.540%	0.111	0.124	-0.013	-11.712%
36 to 40 years old	0.053	0.030	0.023	43.396%	0.053	0.048	0.005	9.434%
41 years old or more	-	-	-	-	-	-	-	-
Migration background	0.194	0.114	0.080	41.237%	0.194	0.185	0.009	4.639%
Educational level achieved:								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.105	0.018	0.087	82.857%	0.105	0.094	0.011	10.476%
High school diploma	-	-	-	-	-	-	-	-
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.371	0.647	-0.276	-74.393%	0.371	0.588	-0.217	-58.491%
<i>Maternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.073	0.036	0.037	50.685%	0.073	0.066	0.007	9.589%
21 to 25 years old	0.247	0.242	0.005	2.024%	0.247	0.249	-0.002	-0.810%
26 to 30 years old	0.189	0.250	-0.061	-32.275%	0.189	0.210	-0.021	-11.112%
31 to 35 years old	0.047	0.026	0.021	44.681%	0.047	0.043	0.004	8.511%
36 to 40 years old	-	-	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-	-	-
Migration background	0.197	0.118	0.079	40.101%	0.197	0.189	0.008	4.061%
Educational level achieved:								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.116	0.060	0.056	48.276%	0.116	0.109	0.007	6.034%
High school diploma	0.190	0.145	0.045	23.684%	0.190	0.186	0.004	2.105%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.727	0.795	-0.068	-9.353%	0.727	0.752	-0.025	-3.439%
<i>Household characteristics</i>								
HH income quintile:								
Missing	-	-	-	-	-	-	-	-
1st quintile	0.185	0.205	-0.020	-10.811%	0.185	0.189	-0.004	-2.162%
2nd quintile	0.189	0.205	-0.016	-8.466%	0.189	0.194	-0.005	-2.645%
3rd quintile	0.239	0.180	0.059	24.686%	0.239	0.228	0.011	4.602%
4th quintile	0.019	0.043	-0.024	-126.316%	0.019	0.027	-0.008	-42.105%
5th quintile	-	-	-	-	-	-	-	-

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: (-) means that the variable was omitted from the matching because of perfect collinearity. The various länder are grouped in four macro-regions. 'North Germany' includes Schleswig-Holstein, Hamburg, Mecklenburg-Vorpommern, Niedersachsen and Bremen. 'South Germany' includes Bayern and Baden-Wuerttemberg. 'Central-West Germany' includes Nordrhein-Westfalen, Hessen, Rheinland-Pfalz and Saarland. Finally, 'Central-East Germany' includes Berlin, Brandenburg, Sachsen, Sachsen-Anhalt and Thuringen.

Table A.5: Pre- and post-matching bias on the mean (treatment 11 to 15 years old)

Variables	Pre-Matching				Post-Matching			
	Variance Treated	Variance Controls	Difference	Bias %	Variance Treated	Variance Controls	Difference	Bias %
<i>Individual characteristics</i>								
Age	23.460	23.135	0.325	1.385%	23.460	23.341	0.119	0.507%
Sex (1 if female)	0.608	0.481	0.127	20.895%	0.608	0.560	0.048	7.971%
Migration background	0.380	0.132	0.247	65.143%	0.380	0.326	0.054	14.265%
Firstborn:								
Missing	0.139	0.103	0.037	26.410%	0.139	0.126	0.014	9.992%
Yes	0.386	0.378	0.007	1.927%	0.386	0.376	0.010	2.584%
Average number of siblings	1.030	1.207	-0.178	-17.267%	1.030	1.080	-0.051	-4.907%
Region where born								
Missing	-	-	-	-	-	-	-	-
North Germany	0.000	0.048	-0.048	-	0.000	0.019	-0.019	-
South Germany	0.068	0.106	-0.038	-55.405%	0.068	0.075	-0.007	-9.687%
Central-West Germany	0.220	0.135	0.085	38.591%	0.220	0.201	0.018	8.356%
Central-East Germany	0.623	0.690	-0.066	-10.653%	0.623	0.630	-0.007	-1.128%
<i>Paternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	-	-	-	-	-	-	-	-
21 to 25 years old	0.255	0.316	-0.061	-23.917%	0.255	0.270	-0.015	-5.936%
26 to 30 years old	0.151	0.322	-0.171	-112.681%	0.151	0.201	-0.049	-32.645%
31 to 35 years old	0.053	0.053	0.001	0.943%	0.053	0.050	0.004	7.213%
36 to 40 years old	0.350	0.129	0.222	63.273%	0.350	0.301	0.049	13.967%
41 years old or more	-	-	-	-	-	-	-	-
Migration background	0.128	0.023	0.105	82.339%	0.128	0.104	0.023	18.242%
Educational level achieved								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.843	0.974	-0.131	-15.523%	0.843	0.871	-0.028	-3.347%
High school diploma	0.409	0.289	0.121	29.446%	0.409	0.369	0.041	10.004%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	0.976	0.408	0.567	58.141%	0.976	0.852	0.123	12.645%
<i>Maternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.427	0.291	0.136	31.812%	0.427	0.385	0.042	9.870%
21 to 25 years old	0.282	0.515	-0.233	-82.604%	0.282	0.347	-0.065	-23.053%
26 to 30 years old	0.231	0.161	0.071	30.470%	0.231	0.217	0.014	6.261%
31 to 35 years old	0.018	0.011	0.007	39.465%	0.018	0.016	0.002	12.841%
36 to 40 years old	-	-	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-	-	-
Migration background	-	-	-	-	-	-	-	-
Educational level achieved								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.662	0.913	-0.252	-38.017%	0.662	0.715	-0.053	-8.009%
High school diploma	0.062	0.010	0.052	83.294%	0.062	0.051	0.012	18.657%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	1.580	0.659	0.922	58.326%	1.580	1.347	0.233	14.746%
<i>Household characteristics</i>								
HH income quintile:								
Missing	-	-	-	-	-	-	-	-
1st quintile	0.261	0.305	-0.044	-16.692%	0.261	0.261	0.000	0.152%
2nd quintile	0.178	0.269	-0.091	-51.337%	0.178	0.204	-0.026	-14.512%
3rd quintile	0.217	0.210	0.007	3.035%	0.217	0.221	-0.004	-2.076%
4th quintile	0.000	0.061	-0.061	-	0.000	0.020	-0.020	-
5th quintile	-	-	-	-	-	-	-	-

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: (-) means that the variable was omitted from the matching because of perfect collinearity. The various länder are grouped in four macro-regions. 'North Germany' includes Schleswig-Holstein, Hamburg, Mecklenburg-Vorpommern, Niedersachsen and Bremen. 'South Germany' includes Bayern and Baden-Wuerttemberg. 'Central-West Germany' includes Nordrhein-Westfalen, Hessen, Rheinland-Pfalz and Saarland. Finally, 'Central-East Germany' includes Berlin, Brandenburg, Sachsen, Sachsen-Anhalt and Thuringen.

Table A.6: Pre- and post-matching bias on the variance (treatment 11 to 15 years old)

Variables	Pre-Matching				Post-Matching			
	Variance Treated	Variance Controls	Difference	Bias %	Variance Treated	Variance Controls	Difference	Bias %
<i>Individual characteristics</i>								
Age	14.130	14.587	-0.457	-3.236%	14.130	14.247	-0.117	-0.831%
Sex (1 if female)	0.239	0.250	-0.011	-4.477%	0.239	0.246	-0.007	-3.127%
Migration background	0.236	0.115	0.121	51.375%	0.236	0.220	0.017	7.040%
Firstborn:								
Missing	0.120	0.092	0.028	23.478%	0.120	0.110	0.011	8.795%
Yes	0.238	0.235	0.002	1.023%	0.238	0.235	0.003	1.285%
Average number of siblings	0.636	0.860	-0.224	-35.244%	0.636	0.750	-0.114	-17.888%
Region where born								
Missing	-	-	-	-	-	-	-	-
North Germany	0.000	0.046	-0.046	-	0.000	0.019	-0.019	-
South Germany	0.064	0.095	-0.031	-48.674%	0.064	0.069	-0.005	-8.599%
Central-West Germany	0.172	0.117	0.055	32.117%	0.172	0.161	0.011	6.468%
Central-East Germany	0.236	0.214	0.021	9.098%	0.236	0.233	0.002	1.041%
<i>Paternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	-	-	-	-	-	-	-	-
21 to 25 years old	0.191	0.216	-0.026	-13.439%	0.191	0.197	-0.007	-3.486%
26 to 30 years old	0.129	0.218	-0.089	-69.463%	0.129	0.160	-0.032	-24.568%
31 to 35 years old	0.051	0.050	0.001	1.172%	0.051	0.047	0.004	7.100%
36 to 40 years old	0.228	0.112	0.116	50.892%	0.228	0.211	0.018	7.756%
41 years old or more	-	-	-	-	-	-	-	-
Migration background	0.112	0.022	0.090	80.268%	0.112	0.093	0.018	16.300%
Educational level achieved								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.133	0.026	0.107	80.623%	0.133	0.112	0.021	15.431%
High school diploma	0.243	0.205	0.037	15.281%	0.243	0.233	0.010	4.035%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	2.374	1.439	0.934	39.365%	2.374	2.406	-0.032	-1.346%
<i>Maternal characteristics</i>								
Age when child was born:								
Missing	-	-	-	-	-	-	-	-
20 years old or less	0.245	0.206	0.039	15.868%	0.245	0.237	0.009	3.508%
21 to 25 years old	0.203	0.250	-0.047	-23.040%	0.203	0.227	-0.024	-11.599%
26 to 30 years old	0.178	0.135	0.043	24.305%	0.178	0.170	0.009	4.765%
31 to 35 years old	0.018	0.011	0.007	39.206%	0.018	0.015	0.002	12.887%
36 to 40 years old	-	-	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-	-	-
Migration background	-	-	-	-	-	-	-	-
Educational level achieved								
Missing	-	-	-	-	-	-	-	-
Less than HS	0.225	0.079	0.145	64.723%	0.225	0.204	0.021	9.172%
High school diploma	0.059	0.010	0.048	82.419%	0.059	0.048	0.010	17.882%
More than HS	-	-	-	-	-	-	-	-
Cumulative unemployment experience	3.533	1.935	1.598	45.220%	3.533	3.287	0.246	6.971%
<i>Household characteristics</i>								
HH income quintile:								
Missing	-	-	-	-	-	-	-	-
1st quintile	0.194	0.212	-0.018	-9.496%	0.194	0.193	0.001	0.382%
2nd quintile	0.147	0.197	-0.050	-34.126%	0.147	0.162	-0.016	-10.597%
3rd quintile	0.170	0.166	0.004	2.499%	0.170	0.172	-0.002	-1.201%
4th quintile	0.000	0.057	-0.057	-	0.000	0.020	-0.020	-
5th quintile	-	-	-	-	-	-	-	-

Source: German Socio-Economic Panel version 36 (SOEP-Core v36).

Notes: (-) means that the variable was omitted from the matching because of perfect collinearity. The various länder are grouped in four macro-regions. 'North Germany' includes Schleswig-Holstein, Hamburg, Mecklenburg-Vorpommern, Niedersachsen and Bremen. 'South Germany' includes Bayern and Baden-Wuerttemberg. 'Central-West Germany' includes Nordrhein-Westfalen, Hessen, Rheinland-Pfalz and Saarland. Finally, 'Central-East Germany' includes Berlin, Brandenburg, Sachsen, Sachsen-Anhalt and Thuringen.

B Items from the SOEP SF-12v2 questionnaire

Table B.1: SOEP SF-12v2 questionnaire

	Bad	Poor	Satisfactory	Good	Very good
(-) How would you describe your current health?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
	Greatly	Often	Sometimes	Almost never	Never
(-) How often in the last four weeks did you feel rushed or under time pressure?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
(-) How often in the last four weeks did you feel in low spirits and melancholy?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
(-) How often in the last four weeks, due to psychological or emotional problems, did you achieve less in your work or everyday activities than you actually intended?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
(-) How often in the last four weeks, due to psychological or emotional problems, did you perform your work or everyday activities less carefully than usual?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
(-) How often in the last four weeks, due to health or psychological problems, have you been restricted in terms of your social contact to for example friends, acquaintances or relatives?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
(-) How often in the last four weeks did you suffer from severe physical pain?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
(-) How often in the last four weeks, due to health problems of a physical nature, have you been restricted in the type of tasks you can perform in your work or everyday activities?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
(-) How often in the last four weeks, due to health problems of a physical nature, did you achieve less in your work or everyday activities than you actually intended?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
	Never	Almost never	Sometimes	Often	Greatly
(-) How often in the last four weeks did you feel calm and balanced?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
(-) How often in the last four weeks did you feel full of energy?	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
	A lot		A little		Not at all
(-) If you have to climb stairs, i.e. walk up several floors: Does your state of health restrict you a lot, a little or not at all?	<input type="checkbox"/>		<input type="checkbox"/>		<input type="checkbox"/>
(-) And what about other strenuous activities in everyday life, e.g. when you have to lift something heavy or need to be mobile: Does your state of health restrict you a lot, a little or not at all?	<input type="checkbox"/>		<input type="checkbox"/>		<input type="checkbox"/>

C Summary statistics tables

Table C.1: Summary statistics for the sample with treatment at 0 to 5 years of age

Variables	Treated (<i>n</i> = 33)				Controls (<i>n</i> = 539)			
	Mean	Standard deviation	Minimum value	Maximum value	Mean	Standard deviation	Minimum value	Maximum value
<i>Dependent variables</i>								
MCS	48.804	7.166	31.208	62.042	50.070	9.095	10.226	67.194
PCS	53.237	6.407	36.245	63.000	55.857	5.795	29.599	68.502
<i>Control variables</i>								
<i>Individual characteristics</i>								
Age	22.000	3.269	18.000	30.000	21.399	3.144	18.000	30.000
Sex (1 if female)	0.545	0.506	0.000	1.000	0.506	0.500	0.000	1.000
Migration background	0.000	0.000	0.000	0.000	0.011	0.105	0.000	1.000
Firstborn:								
Missing	0.121	0.331	0.000	1.000	0.132	0.339	0.000	1.000
Yes	0.394	0.496	0.000	1.000	0.399	0.490	0.000	1.000
Average number of siblings	1.030	0.810	0.000	2.000	0.835	0.688	0.000	4.000
Länder where the child was born:								
Schleswig-Holstein	0.000	0.000	0.000	0.000	0.022	0.148	0.000	1.000
Hamburg	0.000	0.000	0.000	0.000	0.002	0.043	0.000	1.000
Niedersachsen	0.000	0.000	0.000	0.000	0.052	0.222	0.000	1.000
Bremen	0.000	0.000	0.000	0.000	0.004	0.061	0.000	1.000
Nordrhein-Westfalen	0.000	0.000	0.000	0.000	0.109	0.313	0.000	1.000
Hessen	0.000	0.000	0.000	0.000	0.063	0.243	0.000	1.000
Rheinland-Pfalz	0.000	0.000	0.000	0.000	0.069	0.253	0.000	1.000
Baden-Wuerttemberg	0.000	0.000	0.000	0.000	0.106	0.308	0.000	1.000
Bayern	0.000	0.000	0.000	0.000	0.087	0.282	0.000	1.000
Saarland	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Berlin	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Brandenburg	0.121	0.331	0.000	1.000	0.015	0.121	0.000	1.000
Mecklenburg-Vorpommern	0.000	0.000	0.000	0.000	0.015	0.121	0.000	1.000
Sachsen	0.152	0.364	0.000	1.000	0.091	0.288	0.000	1.000
Sachsen-Anhalt	0.000	0.000	0.000	0.000	0.007	0.086	0.000	1.000
Thueringen	0.061	0.242	0.000	1.000	0.020	0.142	0.000	1.000
Deutschland	0.667	0.479	0.000	1.000	0.338	0.473	0.000	1.000
Year of birth:								
1984	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
1985	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
1986	0.212	0.415	0.000	1.000	0.152	0.359	0.000	1.000
1987	0.152	0.364	0.000	1.000	0.058	0.233	0.000	1.000
1988	0.182	0.392	0.000	1.000	0.180	0.385	0.000	1.000
1989	0.000	0.000	0.000	0.000	0.087	0.282	0.000	1.000
1990	0.242	0.435	0.000	1.000	0.167	0.373	0.000	1.000
1991	0.152	0.364	0.000	1.000	0.052	0.222	0.000	1.000
1992	0.061	0.242	0.000	1.000	0.071	0.256	0.000	1.000
1993	0.000	0.000	0.000	0.000	0.045	0.206	0.000	1.000
1994	0.000	0.000	0.000	0.000	0.015	0.121	0.000	1.000
1995	0.000	0.000	0.000	0.000	0.037	0.189	0.000	1.000
1996	0.000	0.000	0.000	0.000	0.056	0.229	0.000	1.000
1997	0.000	0.000	0.000	0.000	0.019	0.135	0.000	1.000
1998	0.000	0.000	0.000	0.000	0.033	0.180	0.000	1.000
1999	0.000	0.000	0.000	0.000	0.019	0.135	0.000	1.000
2000	0.000	0.000	0.000	0.000	0.011	0.105	0.000	1.000
Survey year:								
2002	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
2004	0.030	0.174	0.000	1.000	0.043	0.202	0.000	1.000
2006	0.121	0.331	0.000	1.000	0.109	0.313	0.000	1.000
2008	0.152	0.364	0.000	1.000	0.143	0.350	0.000	1.000
2010	0.242	0.435	0.000	1.000	0.145	0.352	0.000	1.000
2012	0.212	0.415	0.000	1.000	0.148	0.356	0.000	1.000
2014	0.152	0.364	0.000	1.000	0.145	0.352	0.000	1.000
2016	0.091	0.292	0.000	1.000	0.141	0.348	0.000	1.000
2018	0.000	0.000	0.000	0.000	0.126	0.332	0.000	1.000
<i>Paternal characteristics</i>								
Age when the child was born:								

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Table C.1: Continued from previous page

Variables	Treated ($n = 33$)				Controls ($n = 539$)			
	Mean	Standard deviation	Minimum value	Maximum value	Mean	Standard deviation	Minimum value	Maximum value
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
20 years old or less	0.182	0.392	0.000	1.000	0.017	0.128	0.000	1.000
21 to 25 years old	0.333	0.479	0.000	1.000	0.219	0.414	0.000	1.000
26 to 30 years old	0.333	0.479	0.000	1.000	0.330	0.471	0.000	1.000
31 to 35 years old	0.030	0.174	0.000	1.000	0.427	0.495	0.000	1.000
36 to 40 years old	0.121	0.331	0.000	1.000	0.007	0.086	0.000	1.000
41 years old or more	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Migration background	0.000	0.000	0.000	0.000	0.011	0.105	0.000	1.000
Educational level achieved:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Less than HS	0.000	0.000	0.000	0.000	0.006	0.074	0.000	1.000
High school diploma	1.000	0.000	1.000	1.000	0.866	0.341	0.000	1.000
More than HS	0.000	0.000	0.000	0.000	0.128	0.334	0.000	1.000
Cumulative unemployment spell	0.379	0.589	0.000	1.900	0.168	0.415	0.000	2.700
<i>Maternal characteristics</i>								
Age when the child was born:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
20 years old or less	0.303	0.467	0.000	1.000	0.050	0.218	0.000	1.000
21 to 25 years old	0.182	0.392	0.000	1.000	0.404	0.491	0.000	1.000
26 to 30 years old	0.394	0.496	0.000	1.000	0.538	0.499	0.000	1.000
31 to 35 years old	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
36 to 40 years old	0.121	0.331	0.000	1.000	0.007	0.086	0.000	1.000
41 years old or more	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Migration background	0.000	0.000	0.000	0.000	0.011	0.105	0.000	1.000
Educational level achieved:								
Missing	0.000	0.000	0.000	0.000	0.011	0.105	0.000	1.000
Less than HS	0.515	0.508	0.000	1.000	0.735	0.442	0.000	1.000
High school diploma	0.485	0.508	0.000	1.000	0.254	0.436	0.000	1.000
More than HS	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Cumulative unemployment spell	0.276	0.353	0.000	0.900	0.345	0.646	0.000	4.200
<i>Household characteristics</i>								
HH income quintile:								
Missing	0.212	0.415	0.000	1.000	0.065	0.247	0.000	1.000
1st quintile	0.515	0.508	0.000	1.000	0.258	0.438	0.000	1.000
2nd quintile	0.000	0.000	0.000	0.000	0.193	0.395	0.000	1.000
3rd quintile	0.273	0.452	0.000	1.000	0.243	0.429	0.000	1.000
4th quintile	0.000	0.000	0.000	0.000	0.180	0.385	0.000	1.000
5th quintile	0.000	0.000	0.000	0.000	0.061	0.240	0.000	1.000
<i>Contextual characteristics</i>								
Local average unemployment rate	8.810	2.170	6.950	13.329	8.665	2.669	6.047	20.068

Source: German Socio-Economic Panel version 36 (SOEP-Core v36)

Table C.2: Summary statistics for the sample with treatment at 6 to 10 years of age

Variables	Treated ($n = 84$)				Controls ($n = 1251$)			
	Mean	Standard deviation	Minimum value	Maximum value	Mean	Standard deviation	Minimum value	Maximum value
<i>Dependent variables</i>								
MCS	49.100	8.653	28.459	66.461	49.010	9.477	7.394	67.194
PCS	55.161	5.977	34.821	62.554	55.663	6.274	25.896	71.870
<i>Control variables</i>								
<i>Individual characteristics</i>								
Age	21.107	2.892	18.000	30.000	21.498	3.162	18.000	31.000
Sex (1 if female)	0.714	0.454	0.000	1.000	0.519	0.500	0.000	1.000
Migration background	0.238	0.428	0.000	1.000	0.108	0.310	0.000	1.000
Firstborn:								
Missing	0.071	0.259	0.000	1.000	0.086	0.280	0.000	1.000
Yes	0.262	0.442	0.000	1.000	0.417	0.493	0.000	1.000
Average number of siblings	1.321	0.731	0.000	3.000	1.243	0.862	0.000	7.000
Länder where the child was born:								
Schleswig-Holsterin	0.000	0.000	0.000	0.000	0.022	0.145	0.000	1.000
Hamburg	0.000	0.000	0.000	0.000	0.008	0.089	0.000	1.000

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Table C.2: Continued from previous page

Variables	Treated (<i>n</i> = 84)				Controls (<i>n</i> = 1251)			
	Mean	Standard deviation	Minimum value	Maximum value	Mean	Standard deviation	Minimum value	Maximum value
Niedersachsen	0.000	0.000	0.000	0.000	0.046	0.209	0.000	1.000
Bremen	0.000	0.000	0.000	0.000	0.004	0.063	0.000	1.000
Nordrhein-Westfalen	0.060	0.238	0.000	1.000	0.137	0.344	0.000	1.000
Hessen	0.024	0.153	0.000	1.000	0.039	0.194	0.000	1.000
Rheinland-Pfalz	0.012	0.109	0.000	1.000	0.038	0.190	0.000	1.000
Baden-Wuerttemberg	0.048	0.214	0.000	1.000	0.106	0.308	0.000	1.000
Bayern	0.000	0.000	0.000	0.000	0.083	0.276	0.000	1.000
Saarland	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Berlin	0.000	0.000	0.000	0.000	0.007	0.085	0.000	1.000
Brandenburg	0.000	0.000	0.000	0.000	0.006	0.075	0.000	1.000
Mecklenburg-Vorpommern	0.000	0.000	0.000	0.000	0.007	0.085	0.000	1.000
Sachsen	0.000	0.000	0.000	0.000	0.041	0.198	0.000	1.000
Sachsen-Anhalt	0.012	0.109	0.000	1.000	0.000	0.000	0.000	0.000
Thuringen	0.000	0.000	0.000	0.000	0.018	0.131	0.000	1.000
Deutschland	0.845	0.364	0.000	1.000	0.440	0.497	0.000	1.000
Year of birth:								
1984	0.226	0.421	0.000	1.000	0.092	0.289	0.000	1.000
1985	0.060	0.238	0.000	1.000	0.106	0.308	0.000	1.000
1986	0.286	0.454	0.000	1.000	0.090	0.287	0.000	1.000
1987	0.083	0.278	0.000	1.000	0.058	0.235	0.000	1.000
1988	0.119	0.326	0.000	1.000	0.095	0.294	0.000	1.000
1989	0.119	0.326	0.000	1.000	0.042	0.202	0.000	1.000
1990	0.012	0.109	0.000	1.000	0.114	0.317	0.000	1.000
1991	0.000	0.000	0.000	0.000	0.067	0.250	0.000	1.000
1992	0.012	0.109	0.000	1.000	0.105	0.306	0.000	1.000
1993	0.012	0.109	0.000	1.000	0.040	0.196	0.000	1.000
1994	0.012	0.109	0.000	1.000	0.055	0.228	0.000	1.000
1995	0.012	0.109	0.000	1.000	0.033	0.178	0.000	1.000
1996	0.000	0.000	0.000	0.000	0.043	0.203	0.000	1.000
1997	0.000	0.000	0.000	0.000	0.015	0.122	0.000	1.000
1998	0.036	0.187	0.000	1.000	0.024	0.153	0.000	1.000
1999	0.012	0.109	0.000	1.000	0.012	0.109	0.000	1.000
2000	0.000	0.000	0.000	0.000	0.008	0.089	0.000	1.000
Survey year:								
2002	0.060	0.238	0.000	1.000	0.023	0.151	0.000	1.000
2004	0.143	0.352	0.000	1.000	0.070	0.254	0.000	1.000
2006	0.202	0.404	0.000	1.000	0.102	0.303	0.000	1.000
2008	0.214	0.413	0.000	1.000	0.125	0.331	0.000	1.000
2010	0.167	0.375	0.000	1.000	0.153	0.360	0.000	1.000
2012	0.107	0.311	0.000	1.000	0.157	0.364	0.000	1.000
2014	0.036	0.187	0.000	1.000	0.148	0.355	0.000	1.000
2016	0.048	0.214	0.000	1.000	0.125	0.331	0.000	1.000
2018	0.024	0.153	0.000	1.000	0.097	0.296	0.000	1.000
Paternal characteristics								
Age when child was born:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
20 years old or less	0.012	0.109	0.000	1.000	0.007	0.085	0.000	1.000
21 to 25 years old	0.452	0.501	0.000	1.000	0.222	0.416	0.000	1.000
26 to 30 years old	0.357	0.482	0.000	1.000	0.521	0.500	0.000	1.000
31 to 35 years old	0.119	0.326	0.000	1.000	0.213	0.409	0.000	1.000
36 to 40 years old	0.060	0.238	0.000	1.000	0.033	0.178	0.000	1.000
41 years old or more	0.000	0.000	0.000	0.000	0.004	0.063	0.000	1.000
Migration background	0.226	0.421	0.000	1.000	0.102	0.302	0.000	1.000
Educational level achieved:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Less than HS	0.119	0.326	0.000	1.000	0.005	0.069	0.000	1.000
High school diploma	0.881	0.326	0.000	1.000	0.995	0.069	0.000	1.000
More than HS	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Cumulative unemployment spell	0.363	0.426	0.000	1.700	0.289	0.815	0.000	7.800
Maternal characteristics								
Age when child was born:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
20 years old or less	0.071	0.259	0.000	1.000	0.031	0.174	0.000	1.000
21 to 25 years old	0.488	0.503	0.000	1.000	0.371	0.483	0.000	1.000
26 to 30 years old	0.298	0.460	0.000	1.000	0.552	0.497	0.000	1.000

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Table C.2: Continued from previous page

Variables	Treated (<i>n</i> = 84)				Controls (<i>n</i> = 1251)			
	Mean	Standard deviation	Minimum value	Maximum value	Mean	Standard deviation	Minimum value	Maximum value
31 to 35 years old	0.060	0.238	0.000	1.000	0.011	0.105	0.000	1.000
36 to 40 years old	0.083	0.278	0.000	1.000	0.030	0.172	0.000	1.000
41 years old or more	0.000	0.000	0.000	0.000	0.004	0.063	0.000	1.000
Migration background	0.238	0.428	0.000	1.000	0.110	0.313	0.000	1.000
Educational level achieved:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Less than HS	0.071	0.259	0.000	1.000	0.055	0.228	0.000	1.000
High school diploma	0.762	0.428	0.000	1.000	0.836	0.370	0.000	1.000
More than HS	0.167	0.375	0.000	1.000	0.109	0.311	0.000	1.000
Cumulative unemployment spell	0.719	0.729	0.000	4.000	0.517	0.872	0.000	7.700
<i>Household characteristics</i>								
HH income quintile:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
1st quintile	0.274	0.449	0.000	1.000	0.275	0.447	0.000	1.000
2nd quintile	0.179	0.385	0.000	1.000	0.272	0.445	0.000	1.000
3rd quintile	0.440	0.499	0.000	1.000	0.260	0.439	0.000	1.000
4th quintile	0.107	0.311	0.000	1.000	0.147	0.354	0.000	1.000
5th quintile	0.000	0.000	0.000	0.000	0.046	0.210	0.000	1.000
<i>Contextual characteristics</i>								
Local average unemployment rate	10.603	2.870	5.500	19.802	10.036	3.399	5.123	19.112

Source: German Socio-Economic Panel version 36 (SOEP-Core v36)

Table C.3: Summary statistics for the sample with treatment at 11 to 15 years of age

Variables	Treated (<i>n</i> = 69)				Controls (<i>n</i> = 1624)			
	Mean	Standard deviation	Minimum value	Maximum value	Mean	Standard deviation	Minimum value	Maximum value
<i>Dependent variables</i>								
MCS	49.855	8.144	29.174	66.227	49.453	9.557	4.939	72.665
PCS	53.645	7.174	34.973	66.429	55.778	6.465	27.364	73.053
<i>Control variables</i>								
<i>Individual characteristics</i>								
Age	20.609	2.340	18.000	28.000	21.236	3.104	18.000	31.000
Sex (1 if female)	0.667	0.475	0.000	1.000	0.478	0.500	0.000	1.000
Migration background	0.362	0.484	0.000	1.000	0.107	0.309	0.000	1.000
Firstborn:								
Missing	0.145	0.355	0.000	1.000	0.102	0.302	0.000	1.000
Yes	0.420	0.497	0.000	1.000	0.390	0.488	0.000	1.000
Average number of siblings	1.159	0.885	0.000	3.000	1.175	0.911	0.000	8.000
Länder where the child was born:								
Schleswig-Holstein	0.000	0.000	0.000	0.000	0.017	0.130	0.000	1.000
Hamburg	0.000	0.000	0.000	0.000	0.004	0.061	0.000	1.000
Niedersachsen	0.000	0.000	0.000	0.000	0.021	0.143	0.000	1.000
Bremen	0.000	0.000	0.000	0.000	0.002	0.043	0.000	1.000
Nordrhein-Westfalen	0.029	0.169	0.000	1.000	0.126	0.332	0.000	1.000
Hessen	0.058	0.235	0.000	1.000	0.025	0.155	0.000	1.000
Rheinland-Pfalz	0.159	0.369	0.000	1.000	0.021	0.143	0.000	1.000
Baden-Wuerttemberg	0.014	0.120	0.000	1.000	0.069	0.253	0.000	1.000
Bayern	0.029	0.169	0.000	1.000	0.054	0.225	0.000	1.000
Saarland	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Berlin	0.043	0.205	0.000	1.000	0.002	0.043	0.000	1.000
Brandenburg	0.000	0.000	0.000	0.000	0.001	0.035	0.000	1.000
Mecklenburg-Vorpommern	0.000	0.000	0.000	0.000	0.001	0.035	0.000	1.000
Sachsen	0.000	0.000	0.000	0.000	0.016	0.126	0.000	1.000
Sachsen-Anhalt	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Thuringen	0.043	0.205	0.000	1.000	0.002	0.043	0.000	1.000
Deutschland	0.623	0.488	0.000	1.000	0.640	0.480	0.000	1.000
Year of birth:								
1984	0.130	0.339	0.000	1.000	0.057	0.231	0.000	1.000
1985	0.116	0.323	0.000	1.000	0.041	0.198	0.000	1.000
1986	0.072	0.261	0.000	1.000	0.107	0.309	0.000	1.000
1987	0.087	0.284	0.000	1.000	0.100	0.301	0.000	1.000

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Table C.3: Continued from previous page

Variables	Treated (<i>n</i> = 69)				Controls (<i>n</i> = 1624)			
	Mean	Standard deviation	Minimum value	Maximum value	Mean	Standard deviation	Minimum value	Maximum value
1988	0.087	0.284	0.000	1.000	0.123	0.329	0.000	1.000
1989	0.087	0.284	0.000	1.000	0.067	0.249	0.000	1.000
1990	0.043	0.205	0.000	1.000	0.087	0.282	0.000	1.000
1991	0.116	0.323	0.000	1.000	0.049	0.216	0.000	1.000
1992	0.014	0.120	0.000	1.000	0.075	0.263	0.000	1.000
1993	0.014	0.120	0.000	1.000	0.043	0.203	0.000	1.000
1994	0.058	0.235	0.000	1.000	0.050	0.219	0.000	1.000
1995	0.072	0.261	0.000	1.000	0.020	0.141	0.000	1.000
1996	0.043	0.205	0.000	1.000	0.050	0.219	0.000	1.000
1997	0.029	0.169	0.000	1.000	0.032	0.176	0.000	1.000
1998	0.000	0.000	0.000	0.000	0.048	0.214	0.000	1.000
1999	0.029	0.169	0.000	1.000	0.025	0.155	0.000	1.000
2000	0.000	0.000	0.000	0.000	0.026	0.161	0.000	1.000
Survey year:								
2002	0.029	0.169	0.000	1.000	0.014	0.118	0.000	1.000
2004	0.087	0.284	0.000	1.000	0.057	0.231	0.000	1.000
2006	0.130	0.339	0.000	1.000	0.102	0.303	0.000	1.000
2008	0.217	0.415	0.000	1.000	0.127	0.333	0.000	1.000
2010	0.174	0.382	0.000	1.000	0.140	0.348	0.000	1.000
2012	0.130	0.339	0.000	1.000	0.142	0.349	0.000	1.000
2014	0.072	0.261	0.000	1.000	0.138	0.345	0.000	1.000
2016	0.087	0.284	0.000	1.000	0.137	0.344	0.000	1.000
2018	0.072	0.261	0.000	1.000	0.142	0.349	0.000	1.000
<i>Paternal characteristics</i>								
Age when child was born:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
20 years old or less	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
21 to 25 years old	0.449	0.501	0.000	1.000	0.245	0.430	0.000	1.000
26 to 30 years old	0.261	0.442	0.000	1.000	0.318	0.466	0.000	1.000
31 to 35 years old	0.159	0.369	0.000	1.000	0.349	0.477	0.000	1.000
36 to 40 years old	0.043	0.205	0.000	1.000	0.070	0.255	0.000	1.000
41 years old or more	0.087	0.284	0.000	1.000	0.018	0.135	0.000	1.000
Migration background	0.362	0.484	0.000	1.000	0.107	0.309	0.000	1.000
Educational level achieved:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Less than HS	0.145	0.355	0.000	1.000	0.018	0.132	0.000	1.000
High school diploma	0.855	0.355	0.000	1.000	0.982	0.135	0.000	1.000
More than HS	0.000	0.000	0.000	0.000	0.001	0.025	0.000	1.000
Cumulative unemployment spell	0.728	1.312	0.000	7.000	0.308	0.941	0.000	15.000
<i>Maternal characteristics</i>								
Age when child was born:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
20 years old or less	0.087	0.284	0.000	1.000	0.025	0.155	0.000	1.000
21 to 25 years old	0.449	0.501	0.000	1.000	0.242	0.428	0.000	1.000
26 to 30 years old	0.304	0.464	0.000	1.000	0.535	0.499	0.000	1.000
31 to 35 years old	0.130	0.339	0.000	1.000	0.187	0.390	0.000	1.000
36 to 40 years old	0.029	0.169	0.000	1.000	0.012	0.108	0.000	1.000
41 years old or more	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Migration background	0.362	0.484	0.000	1.000	0.106	0.308	0.000	1.000
Educational level achieved:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
Less than HS	0.261	0.442	0.000	1.000	0.060	0.238	0.000	1.000
High school diploma	0.710	0.457	0.000	1.000	0.934	0.248	0.000	1.000
More than HS	0.029	0.169	0.000	1.000	0.006	0.074	0.000	1.000
Cumulative unemployment spell	1.499	1.926	0.000	7.700	0.656	1.405	0.000	10.000
<i>Household characteristics</i>								
HH income quintile:								
Missing	0.000	0.000	0.000	0.000	0.000	0.000	0.000	0.000
1st quintile	0.348	0.480	0.000	1.000	0.146	0.353	0.000	1.000
2nd quintile	0.290	0.457	0.000	1.000	0.323	0.468	0.000	1.000
3rd quintile	0.174	0.382	0.000	1.000	0.283	0.451	0.000	1.000
4th quintile	0.188	0.394	0.000	1.000	0.193	0.395	0.000	1.000
5th quintile	0.000	0.000	0.000	0.000	0.055	0.229	0.000	1.000
<i>Contextual characteristics</i>								
Local average unemployment rate	11.575	4.566	5.642	20.193	9.741	3.801	3.725	20.193

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Table C.3: Continued from previous page

Variables	Treated (<i>n</i> = 69)				Controls (<i>n</i> = 1624)			
	Mean	Standard deviation	Minimum value	Maximum value	Mean	Standard deviation	Minimum value	Maximum value

Source: German Socio-Economic Panel version 36 (SOEP-Core v36)

D Full set of estimation results of the benchmark models

Table D.1: Full set of estimation results with treatment at 0 to 5 years of age

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
Constant	54.386 (100.310)	82.290 (66.186)	28.974 (99.625)	98.256 (67.760)	26.137 (101.590)	94.715 (68.129)
<i>Treatment variable</i>						
Parental unemployment	-2.268 (1.887)	-4.506*** (1.333)	-0.830 (2.158)	-4.612*** (1.535)	-0.818 (2.283)	-4.658*** (1.603)
<i>Individual characteristics</i>						
Age	0.623*** (0.148)	-0.676*** (0.092)	0.636*** (0.148)	-0.670*** (0.092)	0.609*** (0.147)	-0.673*** (0.093)
Sex (1 if female)	-8.218*** (1.727)	-4.297*** (1.210)	-6.583*** (1.913)	-4.528*** (1.357)	-5.656** (2.278)	-4.447*** (1.601)
Migration background	- -	- -	- -	- -	- -	- -
Firstborn:						
Missing	9.075*** (2.625)	0.662 (1.862)	9.419*** (2.671)	0.927 (1.916)	7.709** (3.626)	1.498 (2.557)
Yes	0.433 (2.165)	0.427 (1.531)	-1.240 (2.378)	-0.137 (1.701)	-2.046 (2.766)	-0.319 (1.953)
Average number of siblings	0.588 (1.899)	-1.453 (1.339)	-8.618** (4.055)	-0.285 (2.895)	-9.951** (4.660)	-0.019 (3.283)
<i>Paternal characteristics</i>						
Age when the child was born: (ref. 20 years old or less)						
Missing	- -	- -	- -	- -	- -	- -
21 to 25 years old	-6.080* (3.443)	3.955 (2.431)	0.348 (4.525)	3.449 (3.237)	3.816 (5.378)	5.547 (3.768)
26 to 30 years old	-0.847 (4.332)	3.618 (3.090)	4.416 (5.427)	3.199 (3.904)	8.136 (6.215)	4.872 (4.372)
31 to 35 years old	-4.989 (5.905)	0.281 (4.142)	1.850 (7.277)	0.735 (5.146)	8.868 (7.893)	3.363 (5.511)
36 to 40 years old	-3.603 (8.192)	7.348 (5.814)	-9.573 (8.964)	4.412 (6.431)	-15.273 (11.641)	2.592 (8.160)
41 years old or more	- -	- -	- -	- -	- -	- -
Migration background	- -	- -	- -	- -	- -	- -
Educational level achieved: (ref. Less than HS)						
Missing	- -	- -	- -	- -	- -	- -
High school diploma	-7.581 (13.845)	-16.719* (9.737)	1.468 (15.239)	-11.384 (10.790)	5.115 (16.901)	-9.241 (11.803)
More than HS	-12.582 (12.688)	-15.331* (8.939)	-3.050 (13.934)	-11.547 (9.882)	0.381 (14.836)	-9.550 (10.391)
Cumulative unemployment spell	0.022 (2.285)	2.484 (1.593)	-0.072 (3.507)	4.107* (2.477)	3.018 (3.715)	4.532* (2.589)
Current unemployment spell			-6.270 (7.790)	-4.968 (5.553)	-20.209** (9.516)	-7.855 (6.650)
Expectations on job loss:						
Missing					1.915 (7.214)	2.225 (5.038)

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Table D.1 Continued from previous page

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
Lagged expectations					-1.264 (2.004)	-1.043 (1.406)
<i>Maternal characteristics</i>						
Age when the child was born: (ref. 20 years old or less)						
Missing	-	-	-	-	-	-
21 to 25 years old	1.724 (3.328)	-4.556* (2.367)	-3.877 (3.890)	-4.735* (2.814)	-6.732 (4.140)	-5.400* (2.937)
26 to 30 years old	6.933 (5.471)	5.242 (3.866)	-3.252 (6.919)	3.468 (4.943)	-12.716 (8.892)	1.139 (6.224)
31 to 35 years old	-	-	-	-	-	-
36 to 40 years old	-	-	-	-	-	-
41 years old or more	-	-	-	-	-	-
Migration background	-	12.472 (12.107)	-1.609 (17.504)	-	-	-
Educational level achieved: (ref. Less than HS)						
Missing	-	-	-	-	-	-
High school diploma	-2.827 (13.741)	5.686 (3.941)	-0.087 (6.176)	-7.779 (9.907)	-1.332 (14.206)	-7.373 (9.978)
More than HS	-8.466 (17.106)			-12.004 (12.499)	5.644 (18.596)	-10.088 (13.085)
Cumulative unemployment spell	2.260 (2.626)	-2.774 (1.845)	1.028 (3.825)	-2.131 (2.700)	0.955 (4.080)	-2.469 (2.847)
Current unemployment spell			3.554 (6.786)	-0.686 (4.832)	11.309 (7.939)	2.792 (5.577)
Expectations on job loss						
Missing					-2.475 (5.935)	-1.171 (4.160)
Lagged expectations					-2.845 (1.794)	-0.562 (1.261)
<i>Household characteristics</i>						
HH income quintile: (ref. 1st quintile)						
Missing	-1.243 (9.649)	0.989 (6.761)	-3.554 (9.776)	0.801 (6.916)	-5.451 (9.970)	-0.090 (6.971)
2nd quintile	3.569 (3.132)	1.385 (2.199)	2.286 (3.221)	1.210 (2.284)	1.925 (3.450)	0.969 (2.423)
3rd quintile	4.117* (2.399)	1.321 (1.695)	2.515 (3.043)	2.241 (2.174)	4.701 (3.379)	2.625 (2.375)
4th quintile	-1.378 (3.008)	-0.164 (2.112)	-6.752 (4.414)	0.972 (3.132)	-7.798* (4.639)	0.415 (3.244)
5th quintile	7.950 (6.651)	0.210 (4.667)	0.653 (7.800)	0.938 (5.553)	-7.394 (8.532)	-2.048 (5.991)
Household size			9.801** (3.808)	-1.297 (2.707)	12.528*** (4.286)	-1.106 (3.015)
Home size			-0.035 (0.031)	-0.018 (0.022)	-0.041 (0.039)	-0.008 (0.027)
<i>Contextual characteristics</i>						
Local unemployment rate	-1.096** (0.553)	-0.055 (0.386)	-1.418** (0.627)	-0.253 (0.442)	-1.517** (0.723)	-0.079 (0.501)
Länder dummies	Yes	Yes	Yes	Yes	Yes	Yes
Year of birth dummies	Yes	Yes	Yes	Yes	Yes	Yes
Survey year dummies	Yes	Yes	Yes	Yes	Yes	Yes
R^2	0.495	0.596	0.513	0.601	0.529	0.605
N. indiv.-per-year obs	572	572	572	572	572	572
N. indiv. in sample	199	199	199	199	199	199

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Table D.1 Continued from previous page

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)

Source: German Socio-Economic Panel version 36 (SOEP-Core v36)

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Clustered standard errors at individual level are reported in parentheses.
(-) means that the variable was omitted either because there was perfect collinearity or there were not available observations in the specific subsample.

Table D.2: Full set of estimation results with treatment at 6 to 10 years of age

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
Constant	55.115*** (6.773)	42.610*** (4.748)	52.188*** (8.337)	36.292*** (5.848)	51.358*** (8.356)	35.863*** (5.882)
<i>Treatment variable</i>						
Parental unemployment	0.887 (0.878)	0.705 (0.620)	-0.442 (1.011)	0.395 (0.715)	-0.882 (1.033)	0.871 (0.733)
<i>Individual characteristics</i>						
Age	0.014 (0.097)	0.145** (0.067)	0.018 (0.098)	0.146** (0.067)	0.019 (0.097)	0.147** (0.066)
Sex (1 if female)	-2.375*** (0.901)	-1.024 (0.637)	-2.416*** (0.915)	-1.239* (0.646)	-2.828*** (0.934)	-1.113* (0.662)
Migration background	-4.807 (9.440)	-1.818 (6.618)	-5.919 (9.452)	-2.443 (6.627)	-6.057 (9.382)	-2.465 (6.600)
Firstborn:						
Missing	2.590 (1.917)	-0.115 (1.353)	3.090 (1.939)	0.241 (1.369)	3.562* (2.015)	-0.033 (1.430)
Yes	0.444 (0.999)	2.273*** (0.706)	0.900 (1.016)	2.409*** (0.718)	0.906 (1.020)	2.268*** (0.724)
Average number of siblings	0.222 (0.795)	-1.004* (0.562)	0.570 (1.268)	-2.322*** (0.895)	0.004 (1.334)	-2.307** (0.946)
<i>Paternal characteristics</i>						
Educational level achieved: (ref. Less than HS)						
Missing	-	-	-	-	-	-
21 to 25 years old	8.234** (4.162)	-0.589 (2.930)	9.119** (4.207)	-0.803 (2.963)	4.692 (4.366)	0.547 (3.090)
26 to 30 years old	8.773** (4.244)	-3.310 (2.987)	10.021** (4.298)	-3.467 (3.027)	4.068 (4.492)	-0.889 (3.178)
31 to 35 years old	12.990*** (4.177)	-2.065 (2.937)	14.048*** (4.215)	-1.891 (2.966)	7.524* (4.456)	1.086 (3.152)
36 to 40 years old	10.117* (5.364)	-2.380 (3.775)	11.393** (5.464)	-2.518 (3.848)	3.946 (5.711)	0.967 (4.040)
41 years old or more	-2.439 (5.936)	1.213 (4.177)	-0.114 (6.056)	2.067 (4.262)	-3.702 (6.300)	-
Migration background	-2.088 (5.583)	3.234 (3.926)	0.214 (5.843)	2.569 (4.112)	4.798 (5.896)	-0.689 (4.167)
Educational level achieved: (ref. Less than HS)						
Missing	-	-	-	-	-	-
High school diploma	-5.110** (2.482)	7.202*** (1.750)	-5.492** (2.502)	6.791*** (1.764)	-1.729 (2.667)	5.404*** (1.888)
More than HS	-	-	-	-	-	-
Cumulative unemployment spell	-0.110 (0.649)	-0.382 (0.459)	-1.224 (0.816)	-0.805 (0.577)	-1.178 (0.824)	-0.604 (0.586)
Current unemployment spell			8.943** (3.645)	2.697 (2.580)	8.775** (3.659)	2.365 (2.601)
Expectations on job loss						
Missing					15.326***	-9.137***

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Table D.3 Continued from previous page

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
Lagged expectations					(3.504)	(2.482)
					1.889**	-1.892***
					(0.744)	(0.528)
<i>Maternal characteristics</i>						
Age when the child was born: (ref. 20 years old or less)						
Missing	-	-	-	-	-	-
	-	-	-	-	-	-
21 to 25 years old	-2.252	0.796	-0.802	0.947	0.827	0.043
	(2.018)	(1.432)	(2.114)	(1.499)	(2.208)	(1.571)
26 to 30 years old	-7.188***	2.190	-5.892**	2.286	-3.196	0.340
	(2.383)	(1.690)	(2.443)	(1.732)	(2.627)	(1.870)
31 to 35 years old	-4.521	11.234***	-2.411	10.306**	4.986	6.021
	(5.842)	(4.115)	(5.998)	(4.222)	(6.293)	(4.445)
36 to 40 years old	-6.624**	0.906	-5.475	0.292	-1.994	-2.432
	(3.278)	(2.326)	(3.419)	(2.425)	(3.572)	(2.542)
41 years old or more	-	-	-	-	-	2.228
	-	-	-	-	-	(4.455)
Migration background	7.724	-0.352	6.528	0.603	2.437	3.437
	(11.092)	(7.780)	(11.181)	(7.845)	(11.158)	(7.856)
Educational level achieved: (ref. Less than HS)						
Missing	-	-	-	-	-	-
	-	-	-	-	-	-
High school diploma	-0.618	0.582	0.468	1.841	1.058	1.825
	(2.137)	(1.507)	(2.368)	(1.672)	(2.389)	(1.694)
More than HS	-1.571	-0.504	-0.520	1.040	-0.080	1.193
	(2.475)	(1.747)	(2.668)	(1.885)	(2.698)	(1.914)
Cumulative unemployment spell	-0.152	-1.290***	-0.427	-1.342***	-1.004*	-1.013**
	(0.495)	(0.350)	(0.572)	(0.404)	(0.596)	(0.422)
Current unemployment spell			3.603*	1.434	3.866*	1.764
			(2.092)	(1.479)	(2.156)	(1.530)
Expectations on job loss						
Missing					-0.985	2.201
					(1.956)	(1.388)
Lagged expectations					-0.677	0.750
					(0.689)	(0.489)
<i>Household characteristics</i>						
HH income quintile: (ref. 1st quintile)						
Missing	-	-	-	-	-	-
	-	-	-	-	-	-
2nd quintile	-1.379	1.765*	-0.259	2.091**	-0.252	2.201**
	(1.298)	(0.918)	(1.366)	(0.965)	(1.375)	(0.976)
3rd quintile	1.640	-0.719	2.769**	-0.522	2.036	0.405
	(1.101)	(0.779)	(1.208)	(0.854)	(1.252)	(0.889)
4th quintile	1.547	0.525	2.820	-0.159	2.647	0.415
	(1.614)	(1.137)	(1.911)	(1.349)	(1.919)	(1.359)
5th quintile	-3.067	2.090	-0.898	0.992	-1.402	1.782
	(3.523)	(2.487)	(3.911)	(2.756)	(3.900)	(2.759)
Household size			-0.250	1.640**	0.179	1.649**
			(1.136)	(0.802)	(1.171)	(0.831)
Home size			-0.003	-0.007	-0.012	-0.004
			(0.016)	(0.011)	(0.016)	(0.012)
<i>Contextual characteristics</i>						
Local unemployment	-0.078	-0.146	-0.167	-0.189	-0.076	-0.235
	(0.206)	(0.145)	(0.213)	(0.150)	(0.212)	(0.150)
R^2	0.265	0.222	0.273	0.227	0.289	0.251
N. indiv.-per-year obs.	1,335	1,335	1,335	1,335	1,335	1,335
N. indiv. in sample	464	464	464	464	464	464

Source: German Socio-Economic Panel version 36 (SOEP-Core v36)

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Clustered standard errors at individual level are reported in parentheses. (-) means that the variable was omitted either because there was perfect collinearity or there were not available observations in the specific subsample.

Table D.3: Full set of estimation results with treatment at 6 to 10

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
Constant	55.115*** (6.773)	42.610*** (4.748)	52.188*** (8.337)	36.292*** (5.848)	51.358*** (8.356)	35.863*** (5.882)
<i>Treatment variable</i>						
Parental unemployment	0.887 (0.878)	0.705 (0.620)	-0.442 (1.011)	0.395 (0.715)	-0.882 (1.033)	0.871 (0.733)
<i>Individual characteristics</i>						
Age	0.014 (0.097)	0.145** (0.067)	0.018 (0.098)	0.146** (0.067)	0.019 (0.097)	0.147** (0.066)
Sex (1 if female)	-2.375*** (0.901)	-1.024 (0.637)	-2.416*** (0.915)	-1.239* (0.646)	-2.828*** (0.934)	-1.113* (0.662)
Migration background	-4.807 (9.440)	-1.818 (6.618)	-5.919 (9.452)	-2.443 (6.627)	-6.057 (9.382)	-2.465 (6.600)
Firstborn:						
Missing	2.590 (1.917)	-0.115 (1.353)	3.090 (1.939)	0.241 (1.369)	3.562* (2.015)	-0.033 (1.430)
Yes	0.444 (0.999)	2.273*** (0.706)	0.900 (1.016)	2.409*** (0.718)	0.906 (1.020)	2.268*** (0.724)
Average number of siblings	0.222 (0.795)	-1.004* (0.562)	0.570 (1.268)	-2.322*** (0.895)	0.004 (1.334)	-2.307** (0.946)
Länder where the child was born: (ref. <i>Nordrhein-Westfalen</i>)						
Schleswig-Holstein	1.151 (6.439)	0.386 (4.516)	0.671 (6.457)	0.351 (4.529)	1.871 (6.431)	-0.284 (4.526)
Hamburg	10.295 (6.355)	0.222 (4.477)	11.892* (6.504)	-1.172 (4.580)	9.996 (6.503)	-0.319 (4.598)
Niedersachsen	5.202 (3.839)	5.448** (2.711)	2.185 (4.036)	4.810* (2.851)	1.839 (4.015)	5.361* (2.849)
Bremen	6.503 (26.152)	8.505 (18.404)	5.154 (26.185)	7.143 (18.428)	5.260 (25.960)	6.300 (18.339)
Hessen	-7.347** (2.900)	-1.071 (2.047)	-9.578*** (3.025)	-1.401 (2.136)	-9.518*** (3.037)	-1.050 (2.155)
Rheinland-Pfalz	0.634 (2.690)	7.303*** (1.900)	0.654 (2.730)	7.636*** (1.926)	3.332 (2.807)	6.542*** (1.990)
Baden-Wuerttemberg	-1.517 (2.296)	1.926 (1.624)	-1.559 (2.408)	1.832 (1.702)	-1.183 (2.425)	2.196 (1.722)
Bayern	0.714 (2.372)	-1.063 (1.677)	-0.082 (2.507)	-0.742 (1.772)	0.359 (2.509)	-0.669 (1.781)
Saarland	- (-)	- (-)	- (-)	- (-)	- (-)	- (-)
Berlin	-0.025 (10.427)	-0.080 (7.383)	-0.363 (10.529)	0.476 (7.454)	0.929 (10.459)	-0.726 (7.438)
Brandenburg	-5.926 (10.489)	4.880 (7.444)	-6.582 (10.523)	5.310 (7.466)	-5.772 (10.446)	4.771 (7.446)
Mecklenburg-Vorpommern	5.147 (17.618)	5.090 (12.455)	4.475 (17.639)	4.561 (12.469)	2.986 (17.482)	5.377 (12.411)
Sachsen	-2.422 (3.379)	2.165 (2.385)	-2.536 (3.430)	1.644 (2.421)	-2.312 (3.477)	2.455 (2.464)
Sachsen-Anhalt	-10.688** (4.439)	-1.944 (3.116)	-14.288*** (4.674)	-2.989 (3.283)	-12.904*** (4.701)	-3.846 (3.314)
Thüringen	-1.659 (5.199)	5.055 (3.671)	-4.235 (5.309)	3.911 (3.749)	-4.431 (5.297)	4.686 (3.756)
Deutschland	0.607 (1.737)	0.711 (1.227)	0.937 (1.788)	0.790 (1.263)	0.371 (1.836)	1.419 (1.303)
Year of birth: (ref. 2000)						
1984	-9.770* (5.004)	4.181 (3.506)	-10.429** (5.044)	4.675 (3.533)	-13.856*** (5.292)	7.622** (3.721)
1985	-8.710* (4.986)	6.030* (3.495)	-8.669* (5.034)	7.004** (3.527)	-12.511** (5.277)	10.327*** (3.712)
1986	-8.804* (5.019)	5.412 (3.518)	-8.797* (5.050)	6.140* (3.539)	-13.420** (5.380)	10.102*** (3.785)
1987	-2.117 (-)	8.910** (-)	-2.803 (-)	9.461*** (-)	-7.435 (-)	12.951*** (-)

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Table D.3 Continued from previous page

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
1988	(5.189) -9.342* (5.093)	(3.637) 7.554** (3.570)	(5.227) -9.519* (5.125)	(3.664) 8.293** (3.591)	(5.581) -13.446** (5.371)	(3.928) 11.572*** (3.778)
1989	-7.992 (5.176)	9.905*** (3.627)	-7.991 (5.222)	10.725*** (3.659)	-12.060** (5.460)	14.392*** (3.841)
1990	-5.615 (5.294)	5.024 (3.715)	-5.398 (5.360)	6.260* (3.760)	-9.445* (5.521)	9.219** (3.887)
1991	-8.632* (5.145)	7.078** (3.608)	-8.608* (5.225)	8.383** (3.663)	-15.324*** (5.495)	12.293*** (3.866)
1992	-4.415 (4.986)	6.977** (3.495)	-4.116 (5.066)	8.339** (3.550)	-6.786 (5.140)	9.964*** (3.614)
1993	-8.312 (5.397)	1.218 (3.784)	-8.668 (5.518)	2.718 (3.866)	-12.761** (5.781)	6.401 (4.066)
1994	-9.884* (5.132)	7.430** (3.598)	-10.025* (5.186)	8.410** (3.634)	-11.995** (5.265)	9.893*** (3.702)
1995	-5.788 (5.266)	7.304** (3.689)	-5.505 (5.433)	9.302** (3.805)	-8.186 (5.570)	10.796*** (3.916)
1996	-4.512 (5.166)	3.498 (3.623)	-4.900 (5.280)	4.898 (3.701)	-8.049 (5.431)	7.378* (3.821)
1997	-11.533* (6.572)	5.447 (4.611)	-11.292* (6.630)	6.778 (4.651)	-15.019** (6.752)	9.460** (4.753)
1998	-5.509 (4.571)	6.662** (3.204)	-6.598 (4.733)	8.099** (3.317)	-8.108* (4.777)	9.416*** (3.358)
1999	-3.236 (5.746)	1.620 (4.021)	-2.980 (5.901)	3.444 (4.129)	-5.736 (5.952)	5.346 (4.176)
Survey year: (ref. 2004)						
2002	2.301* (1.385)	-1.048 (0.945)	2.209 (1.385)	-1.186 (0.946)	2.229 (1.382)	-1.176 (0.944)
2006	4.058*** (0.791)	-1.250** (0.539)	4.124*** (0.790)	-1.230** (0.539)	4.190*** (0.785)	-1.241** (0.536)
2008	2.932*** (0.781)	-0.642 (0.533)	3.015*** (0.780)	-0.617 (0.533)	3.010*** (0.776)	-0.594 (0.530)
2010	1.580** (0.798)	-1.559*** (0.545)	1.620** (0.797)	-1.548*** (0.545)	1.535* (0.793)	-1.465*** (0.542)
2012	1.929** (0.898)	-1.546** (0.613)	1.967** (0.897)	-1.527** (0.613)	2.005** (0.892)	-1.525** (0.610)
2014	2.290** (0.999)	0.021 (0.682)	2.309** (0.998)	0.047 (0.682)	2.319** (0.992)	0.046 (0.678)
2016	3.002*** (0.964)	-2.705*** (0.658)	3.010*** (0.962)	-2.678*** (0.657)	3.041*** (0.957)	-2.655*** (0.654)
2018	- -	- -	- -	- -	- -	- -
Paternal characteristics						
Educational level achieved: (ref. Less than HS)						
Missing	- -	- -	- -	- -	- -	- -
21 to 25 years old	8.234** (4.162)	-0.589 (2.930)	9.119** (4.207)	-0.803 (2.963)	4.692 (4.366)	0.547 (3.090)
26 to 30 years old	8.773** (4.244)	-3.310 (2.987)	10.021** (4.298)	-3.467 (3.027)	4.068 (4.492)	-0.889 (3.178)
31 to 35 years old	12.990*** (4.177)	-2.065 (2.937)	14.048*** (4.215)	-1.891 (2.966)	7.524* (4.456)	1.086 (3.152)
36 to 40 years old	10.117* (5.364)	-2.380 (3.775)	11.393** (5.464)	-2.518 (3.848)	3.946 (5.711)	0.967 (4.040)
41 years old or more	-2.439 (5.936)	1.213 (4.177)	-0.114 (6.056)	2.067 (4.262)	-3.702 (6.300)	- -
Migration background	-2.088 (5.583)	3.234 (3.926)	0.214 (5.843)	2.569 (4.112)	4.798 (5.896)	-0.689 (4.167)
Educational level achieved: (ref. Less than HS)						
Missing	- -	- -	- -	- -	- -	- -
High school diploma	-5.110**	7.202***	-5.492**	6.791***	-1.729	5.404***

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Table D.3 Continued from previous page

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
More than HS	(2.482)	(1.750)	(2.502)	(1.764)	(2.667)	(1.888)
	-	-	-	-	-	-
Cumulative unemployment spell	-0.110	-0.382	-1.224	-0.805	-1.178	-0.604
	(0.649)	(0.459)	(0.816)	(0.577)	(0.824)	(0.586)
Current unemployment spell			8.943**	2.697	8.775**	2.365
			(3.645)	(2.580)	(3.659)	(2.601)
Expectations on job loss						
Missing					15.326***	-9.137***
					(3.504)	(2.482)
Lagged expectations					1.889**	-1.892***
					(0.744)	(0.528)
<i>Maternal characteristics</i>						
<i>Age when the child was born: (ref. 20 years old or less)</i>						
Missing	-	-	-	-	-	-
	-	-	-	-	-	-
21 to 25 years old	-2.252	0.796	-0.802	0.947	0.827	0.043
	(2.018)	(1.432)	(2.114)	(1.499)	(2.208)	(1.571)
26 to 30 years old	-7.188***	2.190	-5.892**	2.286	-3.196	0.340
	(2.383)	(1.690)	(2.443)	(1.732)	(2.627)	(1.870)
31 to 35 years old	-4.521	11.234***	-2.411	10.306**	4.986	6.021
	(5.842)	(4.115)	(5.998)	(4.222)	(6.293)	(4.445)
36 to 40 years old	-6.624**	0.906	-5.475	0.292	-1.994	-2.432
	(3.278)	(2.326)	(3.419)	(2.425)	(3.572)	(2.542)
41 years old or more	-	-	-	-	-	2.228
	-	-	-	-	-	(4.455)
Migration background	7.724	-0.352	6.528	0.603	2.437	3.437
	(11.092)	(7.780)	(11.181)	(7.845)	(11.158)	(7.856)
<i>Educational level achieved: (ref. Less than HS)</i>						
Missing	-	-	-	-	-	-
	-	-	-	-	-	-
High school diploma	-0.618	0.582	0.468	1.841	1.058	1.825
	(2.137)	(1.507)	(2.368)	(1.672)	(2.389)	(1.694)
More than HS	-1.571	-0.504	-0.520	1.040	-0.080	1.193
	(2.475)	(1.747)	(2.668)	(1.885)	(2.698)	(1.914)
Cumulative unemployment spell	-0.152	-1.290***	-0.427	-1.342***	-1.004*	-1.013**
	(0.495)	(0.350)	(0.572)	(0.404)	(0.596)	(0.422)
Current unemployment spell			3.603*	1.434	3.866*	1.764
			(2.092)	(1.479)	(2.156)	(1.530)
Expectations on job loss						
Missing					-0.985	2.201
					(1.956)	(1.388)
Lagged expectations					-0.677	0.750
					(0.689)	(0.489)
<i>Household characteristics</i>						
<i>HH income quintile: (ref. 1st quintile)</i>						
Missing	-	-	-	-	-	-
	-	-	-	-	-	-
2nd quintile	-1.379	1.765*	-0.259	2.091**	-0.252	2.201**
	(1.298)	(0.918)	(1.366)	(0.965)	(1.375)	(0.976)
3rd quintile	1.640	-0.719	2.769**	-0.522	2.036	0.405
	(1.101)	(0.779)	(1.208)	(0.854)	(1.252)	(0.889)
4th quintile	1.547	0.525	2.820	-0.159	2.647	0.415
	(1.614)	(1.137)	(1.911)	(1.349)	(1.919)	(1.359)
5th quintile	-3.067	2.090	-0.898	0.992	-1.402	1.782
	(3.523)	(2.487)	(3.911)	(2.756)	(3.900)	(2.759)
Household size			-0.250	1.640**	0.179	1.649**
			(1.136)	(0.802)	(1.171)	(0.831)
Home size			-0.003	-0.007	-0.012	-0.004
			(0.016)	(0.011)	(0.016)	(0.012)
<i>Contextual characteristics</i>						
Local unemployment	-0.078	-0.146	-0.167	-0.189	-0.076	-0.235

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Table D.3 Continued from previous page

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
	(0.206)	(0.145)	(0.213)	(0.150)	(0.212)	(0.150)
Länder dummies	Yes	Yes	Yes	Yes	Yes	Yes
Year of birth dummies	Yes	Yes	Yes	Yes	Yes	Yes
Survey year dummies	Yes	Yes	Yes	Yes	Yes	Yes
R^2	0.265	0.222	0.273	0.227	0.289	0.251
N. indiv.-per-year obs.	1335	1335	1335	1335	1335	1335
N. indiv. in sample	464	464	464	464	464	464

Source: German Socio-Economic Panel version 36 (SOEP-Core v36)

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Clustered standard errors at individual level are reported in parentheses. (-) means that the variable was omitted either because there was perfect collinearity or there were not available observations in the specific subsample.

Table D.4: Full set of estimation results with treatment at 11 to 15 years of age

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
Constant	42.570*** (4.414)	69.644*** (3.154)	36.620*** (5.485)	71.967*** (3.898)	35.700*** (5.692)	71.449*** (4.024)
<i>Treatment variable</i>						
Parental unemployment	0.739 (0.720)	-1.966*** (0.512)	1.401* (0.812)	-2.252*** (0.575)	1.381 (0.857)	-2.597*** (0.603)
<i>Individual characteristics</i>						
Age	0.246*** (0.081)	-0.294*** (0.059)	0.239*** (0.081)	-0.290*** (0.059)	0.236*** (0.081)	-0.282*** (0.059)
Sex (1 if female)	-3.539*** (0.738)	-1.191** (0.525)	-3.496*** (0.760)	-1.447*** (0.539)	-3.490*** (0.765)	-1.436*** (0.540)
Migration background	5.507 (9.241)	-2.298 (6.550)	3.822 (9.261)	-2.186 (6.542)	4.727 (9.280)	-0.259 (6.525)
Firstborn:						
Missing	-0.612 (1.277)	-1.812** (0.908)	0.265 (1.489)	-2.542** (1.055)	0.219 (1.505)	-2.374** (1.061)
Yes	0.763 (0.936)	-0.520 (0.665)	0.705 (0.962)	-0.816 (0.682)	0.924 (0.967)	-1.006 (0.682)
Average number of siblings	-0.604 (0.574)	-1.433*** (0.408)	-0.557 (0.893)	-0.720 (0.633)	-0.479 (0.936)	-0.600 (0.660)
<i>Paternal characteristics</i>						
Age when the child was born: (ref. 20 years old or less)						
21 to 25 years old	-1.948 (2.143)	2.835 (7.465)	-	-	-	-
26 to 30 years old	-4.920*** (1.725)		-2.913** (1.411)	-0.153 (7.543)	-2.191 (1.461)	2.021 (7.518)
31 to 35 years old	-4.382** (1.809)	-2.930** (1.228)	-1.933 (1.510)	-2.059** (0.998)	-1.178 (1.555)	-3.081*** (1.029)
36 to 40 years old	-3.269 (2.211)	-4.084*** (1.287)	-1.544 (2.234)	-2.661** (1.068)	-1.455 (2.233)	-3.540*** (1.095)
41 years old or more	-	-2.129 (1.573)	0.998 (2.194)	0.042 (1.583)	1.220 (2.214)	0.240 (1.574)
Migration background	-4.816 (10.519)	-1.316 (1.522)	-2.438 (10.663)	-	-	-
Educational level achieved: (ref. Less than HS)						
Missing	-	-	-	-	-	-
High school diploma	1.629 (1.892)		1.493 (1.914)	1.316 (1.555)	1.195 (2.007)	0.770 (1.561)
More than HS	-	-2.923**	-	-3.153**	-2.147	-1.673

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Table D.4 Continued from previous page

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
Cumulative unemployment spell	- 0.032 (0.294)	(1.343) 0.126 (0.209)	- 0.100 (0.419)	(1.354) 0.008 (0.298)	(10.679) -0.095 (0.436)	(1.412) 0.067 (0.308)
Current unemployment spell			-0.745 (2.033)	1.572 (1.442)	-0.002 (2.196)	-0.116 (1.549)
Expectations on job loss:						
Missing					-2.595 (2.391)	4.146** (1.685)
Lagged expectations					-0.525 (0.838)	-0.763 (0.590)
<i>Maternal characteristics</i>						
Age when te child was born: (ref. 20 years old or less)						
Missing	-	-	-	-	-	-
21 to 25 years old	4.674*** (1.746)	2.510** (1.234)	4.854*** (1.751)	2.427** (1.233)	3.375* (1.911)	1.564 (1.341)
26 to 30 years old	7.687*** (2.173)	2.902* (1.537)	7.964*** (2.200)	3.008* (1.551)	5.926** (2.413)	3.030* (1.695)
31 to 35 years old	5.727** (2.461)	3.221* (1.743)	5.780** (2.478)	2.666 (1.750)	4.505* (2.581)	2.244 (1.814)
36 to 40 years old	-4.946 (3.381)	2.305 (2.404)	-4.166 (3.417)	1.658 (2.423)	-5.376 (3.496)	0.763 (2.467)
41 years old or more	-	-	-	-	-	-
Migration background	0.641 (6.238)	0.353 (6.608)	-0.123 (6.511)	0.972 (4.625)	-3.285 (9.345)	-2.211 (4.674)
Educational level achieved: (ref. Less than HS)						
Missing	-	-	-	-	-	-
High school diploma	-0.981 (1.484)	0.253 (1.055)	-1.009 (1.516)	-0.062 (1.075)	-0.763 (1.576)	-0.845 (1.111)
More than HS	0.649 (3.648)	0.559 (2.597)	0.283 (3.827)	1.424 (2.717)	-0.681 (3.910)	-0.388 (2.761)
Cumulative unemployment spell	-0.116 (0.211)	0.860*** (0.150)	-0.262 (0.253)	0.994*** (0.179)	-0.279 (0.266)	0.701*** (0.187)
Current unemployment spell			2.070 (1.512)	-1.527 (1.071)	0.988 (1.611)	-0.037 (1.137)
Expectations on job loss:						
Missing					2.819 (2.016)	-0.306 (1.420)
Lagged expectations					1.544* (0.790)	0.809 (0.556)
<i>Household characteristics</i>						
HH income quintile: (ref. 1st quintile)						
Missing	-	-	-	-	-	-
2nd quintile	1.164 (0.983)	1.133 (0.698)	0.849 (0.999)	1.118 (0.707)	0.880 (1.021)	1.016 (0.719)
3rd quintile	0.052 (1.060)	-1.988*** (0.753)	-0.360 (1.120)	-1.552* (0.793)	-0.608 (1.184)	-1.714** (0.834)
4th quintile	2.107* (1.164)	-0.196 (0.828)	1.190 (1.300)	0.106 (0.922)	0.640 (1.356)	-0.501 (0.957)
5th quintile	2.045 (3.388)	1.123 (2.405)	1.213 (3.425)	1.230 (2.423)	0.746 (3.455)	0.483 (2.433)
Household size			0.081 (0.989)	-1.172* (0.702)	0.395 (1.000)	-1.415** (0.706)
Home size			0.027** (0.012)	0.012 (0.009)	0.026** (0.012)	0.013 (0.009)
<i>Contextual characteristics</i>						
Local unemployment rate	0.030 (0.122)	-0.410*** (0.087)	-0.016 (0.130)	-0.368*** (0.092)	-0.046 (0.132)	-0.342*** (0.093)

(continued on next page)

Table D.4 *Continued from previous page*

	Baseline specification		Not including parental job loss expectations		Including parental job loss expectations	
	MCS	PCS	MCS	PCS	MCS	PCS
	(1a)	(1b)	(2a)	(2b)	(3a)	(3b)
Länder dummies	Yes	Yes	Yes	Yes	Yes	Yes
Year of birth dummies	Yes	Yes	Yes	Yes	Yes	Yes
Survey year dummies	Yes	Yes	Yes	Yes	Yes	Yes
R^2	0.349	0.372	0.353	0.378	0.357	0.405
N. indiv.-per-year obs.	1693	1693	1693	1693	1693	1693
N. indiv. in sample	664	664	664	664	664	664

Source: German Socio-Economic Panel version 36 (SOEP-Core v36)

Notes: * Significant at 10%, ** significant at 5% and *** significant at 1%. Clustered standard errors at individual level are reported in parentheses.
(-) means that the variable was omitted either because there was perfect collinearity or there were not available observations in the specific subsample.