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Cash Transfers and Labor Supply: New Evidence on Impacts and Mechanisms

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Abstract

We study the impact of a national cash transfer program in Vietnam on labor supply using large household surveys and a regression-discontinuity design based on discontinuity in age eligibility. We do not find evidence of a disincentive effect of the cash transfer on labor supply for adults aged 15-64. More importantly, we find robust evidence that the transfer program causes the adults to move from self-employed non-farm work to wage-paying jobs. A likely mechanism is that the transfer program reduces the labor force participation of older people, and they help housework and childcare for younger adults to have wage-paying jobs.

JEL classification: J22; N35; H55.

Keywords: Cash transfer; social security; employment; labor market participation; Vietnam.

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1. Introduction

Cash transfers are a crucial social assistance policy to help individuals or households in need, and most countries have some kind of cash transfer scheme in place. Although there is widespread consensus on the potential positive welfare impacts of cash transfers, there is in many cases a concern that cash transfers can reduce the incentives to work and create a poverty trap for the recipients. Understanding the impact and mechanism of cash transfers on labor supply is of importance for improving the policy design of cash transfers with potentially important impacts on welfare outcomes.

In this study, we examine the effect of cash transfers from the social assistance policy on individual labor supply in Vietnam using the Vietnam Household Living Standard Surveys (henceforth referred to as VHLSS) from 2012 to 2020. The government of Vietnam has provided monthly cash transfers for social protection beneficiaries including infant children, people with disability, and older people in poor households. Since 2011, the cash transfer program has been extended to cover all people aged 80 or more, who do not have contributory pensions. Between 2012 and 2020, the cash transfer program was referred to as cash transfers for social protection beneficiaries according to Decree 136/2013/ND-CP of the government of Vietnam (Government of Vietnam, 2013). In 2020, there were around 3 million people receiving the cash transfer (around 3% of the total population), and the proportion of older people among all the transfer recipients was 55% (MOLISA, 2022). On average, the cash transfer account for around 10% of the total household income (estimated from the Vietnam Household Living Standard Surveys). Using the discontinuity in age eligibility at 80 years old, we apply a regression discontinuity design to assess the impact of the cash transfer from the social assistance policy on the employment behavior of individuals aged 15-64. We do not find evidence that the cash transfer reduces the incentives to work of individuals aged 15-64. However, we find robust evidence that it does have a significant effect on the transition of labor from non-farm self-employment into wage-paying employment. Specifically, the cash transfer received by households reduces the probability of having self-employed non-farm work by 0.13, but increases the probability of having a wage-paying job by 0.14.

We argue that childcare provided by older people is one of the main mechanisms through which the cash transfer program induces labor mobility of individuals from selfemployed non-farm work to wage-paying jobs. We find that the cash transfer reduces the probability of working for people aged from 65 years and above. In Vietnam, 30% of children under 6 years old live with grandparents (according to the 2020 VHLSS), and it is common that grandparents provide childcare for small children. Recent studies show that becoming grandparents substantially reduces the labor force participation of older people (Rupert and Zanella, 2018; Frimmel et al., 2022), and childcare by grandparents significantly increases women's labor supply of (e.g., Maurer-Fazio et al., 2011; Zanella, 2017; Garcia-Moran, E., & Kuehn, 2017; Halim et al., 2021). For our case study, the cash transfer reduces the labor force participation of older people to find wage jobs in the labor market. This argument is consistent with our analysis which shows that the effect of the cash transfer on the probability of having a wage job is found to be higher for women and those having small children.

Our study aims to make two contributions to the literature on the effect of public cash transfers (see the review from Rawlings and Rubio, 2005; Kabeer and Waddington, 2015; Handa et al., 2018; Bastagli et al., 2019). Our first contribution is related to the literature on the response of the labor supply to public transfers. According to traditional economic theory, leisure is a normal good and the wage rate is fixed, and as a result an increase in non-labor income such as cash transfers may reduce the labor supply. There is some evidence that transfers reduce the incentive to work and the labor market participation of recipients (e.g., Lloyd-Sherlock, 2006; Farrington and Slater, 2006; Abel, 2019). However, in low- and middle-income countries, people may face credit constraints. Cash transfers are fungible and can be used for investment and other productive purposes instead of leisure (e.g., Kabeer and Waddington, 2015; Gertler et al., 2012; and Blattman et al., 2014), in which case they do not decrease the labor supply. Disincentive effects of cash transfers on the labor force participation are not found in a number of studies such as Kuhn et al. (2011), Barrientos and Villa (2015), Banerjee et al. (2017), Baird et al. (2018), Ham and Michelson (2018), Jones and Marinescu (2022), and Vera-Cossio (2022).¹ In this study we show that the effect of the cash transfer program in Vietnam differs by age groups. Importantly, the cash transfer is largely targeted at older people, and the amount of cash transfers is not large enough to change the labor supply of younger people. However, for older people, who have low earnings, the cash transfer can compensate for their earnings loss due to not working. Moreover, for some older people, working can increase work-related health problems, and retirement can help them to spend more time on leisure and health care and improve their

¹ For a review of the impact of cash transfers, see Kabeer and Waddington (2015), Parker and Todd (2017), Baird et al. (2018) and Bastagli et al. (2019).

physical and mental health (e.g., Atalay and Barrett, 2014; Zhu, 2016 Gorry et al., 2018; Rose, 2020).

Our second contribution is as noted to show that the cash transfer can promote labor transition from non-farm self-employment to wage-paying jobs for younger people arguably through the channel of childcare. Several studies such as Lindh et al. (1996), Blattman et al. (2014) and Salehi-Isfahani and Mostafavi-Dehzooei (2018) document that cash transfers can be invested in non-farm business, thereby increasing non-farm self-employment of individuals. There are two possible reasons for the difference between these studies: large non-labor income from lottery winnings and inheritances in Lindh et al. (1996); the transfer equal to the average annual income in Blattman et al. (2014); and the transfer equal to 20% of total expenditure in Salehi-Isfahani and Mostafavi-Dehzooei (2018). In contrast, in our study, the transfer accounted for around 10% of the total household income. Second, the cash transfer in our study is targeted at the most disadvantageous groups (infant children, people with disability, and older people without contributory pensions).

Vietnam is a lower middle-income country in Southeast Asia, which has achieved impressive success in economic growth and poverty reduction. Vietnam has implemented numerous poverty reduction and social assistance programs. Impacts of cash transfers are studied in several studies (e.g., Van de Walle, 2004; Van den Berg and Nguyen 2011; Nguyen, 2013; Nguyen, 2021). These studies focus on the effect of cash transfers on household welfare and poverty and most studies find a significant effect of cash transfers on welfare improvement and poverty reduction. Our study differs from these previous studies in Vietnam in two aspects. Firstto our knowledge, the present study is the first attempt to estimate the effect of cash transfers on labor supply in Vietnam. In Vietnam, there is still a large proportion of self-employment, at 36% according to the 2020 VHLSS. Second, compared with previous studies, which used data sets before 2012, our study provides more updated findings on the impact of cash transfers by using recent household surveys from 2012 to 2020.

This paper is structured in six sections. Section 2 presents data sets and descriptive analysis, while Sections 3 and 4 cover the estimation strategy and the empirical results of the impact of the cash transfer on labor supply. Section 5 discusses potential mechanisms through which the cash transfer changes the labor behaviors of individuals, and Section 6 concludes.

2. Data and country context

The VHLSS surveys from 2010 to 2020 were conducted by the General Statistics Office of Vietnam (GSO) with technical support from the World Bank. Each VHLSS covers around 45,000 households from around 3,000 enumeration areas and includes detailed socioeconomic data on households and their members. These VHLSSs are representative at the provincial and regional level. We use the sampling weights, computed by GSO and the World Bank, in the estimations and regressions.

The VHLSSs contain information on households and members living in the households. Household-level data include information on households' assets, production, income, housing condition, and participation in government programs. Individual-level data consist of information on demographics, health, education, and employment. Regarding employment, employment characteristics such as sectors and occupations are measured for

the main job during the past 12 months. The number of working hours and wages are measured for the past 30 days. The VHLSSs contain information on cash transfers according to the Decree 136/2013/NĐ-CP that households received during the past 12 months.²

Vietnam has committed to follow a 'growth with equity' strategy of development. The cash transfer program has provided a regular, monthly, cash transfer for the social protection group since 2000 (Government of Vietnam, 2000). The cash transfer program is targeted at the most vulnerable groups including infants and children below 16 years of age, people with heavy disability, poor older people living alone without any support. Between 2007 and 2011, the coverage of the cash transfer program was extended to older people, 85 and above (Government of Vietnam, 2007). Since 2011, the age threshold for older people to receive the cash transfer has been reduced to 80 (Government of Vietnam, 2011). During the 2012-2020 period, the beneficiaries of the cash transfer program according to Decree 136/2013/NĐ-CP include the following groups (Government of Vietnam, 2013):³

- Infants and children below 16 years old and children with high disability.
- People with HIV disease and people with high disability.
- People from poor households without a spouse and currently raising a child under 16 years old.
- People from 60 years living in poor households and not having any supports.
- People aged 80 and above who do not have contributory pensions.

² In Vietnam, there are annual labor force surveys which contain information on employment. However, we do not use this data set, since it does not contain information on cash transfers received by households.
³ In 2021, Decree 136/2013/NĐ-CP was revised and replaced by Decree 20/2021/NĐ-CP on social support policies for social protection beneficiaries. The main difference between Decree 136/2013/NĐ-CP and Decree 20/2021/NĐ-CP is that the latter provides a larger amount of cash transfers.

In Vietnam, there are other types of cash transfers such as allowances for war invalids. We focus here on the impact of the cash transfer from the social assistance policy under Decree 136/2013/NĐ-CP, which is provided for individuals. The VHLSSs do not collect information of the cash transfer received by individuals, but on the total cash transfers from the social assistance policy that households received during the past 12 months. Thus, it is possible that there are more than one household member receiving the cash transfer.

Figure 1 shows that the proportion of households receiving the cash transfer from the social assistance policy increased substantially between 2010 and 2012, since the age threshold for the cash transfer was reduced from 85 to 80. According MOLISA (2022), 55% of the recipients of the cash transfer in 2020 were older people. That is why as the age threshold was decreased, the proportion of households with the cash transfer increased remarkably. The proportion of transfer-receiving households increased over time between 2012 and 2018, but this rate decreased from 12.6% in 2018 to 10.4% in 2020. A possible reason for this decrease is the change in the sampling frame for the VHLSSs. The VHLSSs from 2010 to 2018 are based on the sampling frame from the 2009 Population and Housing Census. The sampled communes were kept unchanged during this period. On the other hand, the 2019 Population and Housing Census was used as the sampling frame for the 2020 VHLSS.

The amount of the cash transfer increased over time. In 2020, the amount of monthly cash transfer ranges from 270 thousand VND/person (around 12 USD according to the exchange rate in 2020) to 810 thousand VND/person (around 35 USD). The average per capita transfer among receiving households was 2,047 thousand/year in 2010, but remarkably dropped to 1,464 thousand VND in 2012. The main reason is that the cash transfer program

provides a higher transfer amount for people with disability and the cash transfer program was mainly provided for people with disability before 2011. Since 2011, a large proportion of recipients are people from 80 years old, but this group receives a lower amount of cash transfers. In 2020, the per capita cash transfer was 2,087 thousand VND.

[Figure 1 around here]

In this study, we focus on individuals aged 15-64, who account for 95% of the labor force. We examine the effect of the cash transfer program on a number of employment outcomes including working, the number of working hours during the last month, monthly wages, and whether workers have a wage job, self-employed farm and nonfarm work, or a job with social insurance (i.e., a formal job). In the VHLSSs, there is no information to define unemployment, but this rate is low in Vietnam, at around 2% (Dang et al., 2023). Table 1 presents the mean and standard error of the outcome variables for years 2022, 2016 and 2020. For presentation simplicity, we do not report the estimates for 2014 and 2018. The rate of people working is guite stable over time, and equal to 82% in 2020. The proportion of people having a wage-paying job increased from 38% in 2012 to 47% in 2020. During this period, the proportion of people with self-employed non-farm work increased slightly from 11% to 15%, while the proportion of people with self-employed farm work decreased from 32% to 20% during this period. People are more likely to work in the formal sector with the proportion increasing from 14% in 2012 to 22% in 2020. Among working people, the number of working hours during the last month increased from 183 to 193 during the 2012-2020 period. The real monthly wages of wage workers also increased over time.⁴ In addition to the

⁴ Month wages in all years are deflated to the prices in December 2020 using monthly CPIs published by the GSO.

mean wage computed for wage workers, we estimate the mean wage, which is averaged across all people aged 15-64. This wage variable captures not only the increase in wages of wage workers but also the increase in the share of wage workers.

Table 1 shows that individuals in households receiving the cash transfer have a lower working rate and fewer working hours than those in households not receiving the cash transfer. Since the cash transfer is targeted at the poorer and older households, individuals in the recipient households tend to have a lower proportion of having a wage job as well as a formal job than those in non-recipient households. They also have a lower wage than those in households not receiving the cash transfer.

[Table 1 around here]

3. Estimation strategy

In our analysis, we rely on a fuzzy Regression Discontinuity Design (RDD) using the discontinuity in age eligibility of the oldest member in households (for more detailed discussion of RDD, see Imbens and Lemieux, 2008; Lee and Lemieux, 2010; Cattaneo et al., 2019). Moreover, we estimate the effect of households receiving the cash transfer on employment outcomes of household individual members aged 15-64.

As discussed in the previous section, one of the main groups of beneficiaries of the cash transfer program includes people, who are older than 80 years and do not have contributory pensions. As a result, households with a member who is 80 years old or more are more like to receive cash transfers than others. Accordingly, there is a discontinuity in the probability of receiving the cash transfer at the age 80. For example, for the 2020 VHLSS, the proportion of households, whose oldest member is aged 79, and which receive the cash transfer was 19%, but this proportion increased to 52% for households when the oldest

member is of age 80. Panel A of Figure 2 plots the probability of households receiving the cash transfer across the age of the oldest in the households. It shows a large gap in the probability of receiving the cash transfer at age 80 years. In Panel B of Figure 2, we measure the age of the oldest by months instead of years. Panel C of Figure 2 shows a large gap in the probability of individuals aged 15-64 living in households, who received the cash transfer at the age 80 of the oldest member. In Panel B of Figure 2, when the age of the oldest is measured by months, we see a discontinuity in the probability of receiving the transfer at the age 960 months when the oldest changed their age from 79 to 80. Since not only people from the age of 80 years, but also those below that age receive the cash transfer, we apply a fuzzy RDD instead of a sharp RDD.

[Figure 2 around here]

It should be noted that people aged 60 and more living in poor households are also eligible for the cash transfer from the social assistance policy. However, the proportion of poor households is relatively small, around 8% during the 2012-2020 period. Figure A.1 shows that there is not a discontinuity in the probability of receiving the cash transfer at age 60. In Figure A.2, we plot the probability of households receiving the cash transfer at other age thresholds (65, 70, 75, and 85) and do not find a discontinuity in the probability at these thresholds. Thus, we do not use these thresholds to estimate the effect of the cash transfer.

Both parametric and non-parametric methods are used. Non-parametric methods do not impose a functional form on the outcomes and can provide estimates close to the cut-off. However, non-parametric methods are often computationally costly and still rely on a functional form such as a kernel function to weigh the observations around the cut-off. Imbens and Gelman (2014) suggest that the local linear of the running variable should be used instead of higher-order polynomials. In this study, we mainly rely on a parametric regression, though we also use a nonparametric method to check robustness. More specifically, we estimate the effect of cash transfers using a local linear regression developed by Hahn et al. (2001) as follows.

$$y_{i,j} = \beta_0 + D_j \beta_1 + (Age_j - 960) \cdot Z_j \cdot \beta_2 + (Age_j - 960) \cdot \beta_3 + X_{i,j} \beta_4 + \varepsilon_{i,j}, \quad (1)$$

where $y_{i,j}$ is the outcome of labor supply of individual *i* in household *j*. D_j is the treatment variable that indicates that the household receives cash transfers from the social assistance policy. Age_j is age in months of the oldest member in the household. Z_j is equal to $I\{Age_j \ge 960\}$, that is a dummy indicating whether the oldest member is equal to or older than 960 months (or 80 years). The threshold or cutoff is equal to 960. We use age in months of the oldest instead of age in years to allow more continuity in the running variable around the cutoff point.⁵ The interaction between the treatment variable and age allows for the effect of age, which differs from the treatment status, and the local average treatment effect τ_{FRD} is estimated by the coefficient of the treatment variable, β_1 . We estimate model (1) using a 2SLS regression, in which Z_j is used as an instrument for D_j .

We control for exogenous individual-level variables $X_{i,j}$ including gender, age, age squared, urban and year dummies to improve the efficiency (Imbens and Lemieux 2008; Calonico et al., 2018). However, we use a limited exogenous control variable, since the control variables should be exogenous and not affected by the cash transfer (Angrist and

⁵ In the VHLSSs, there are questions on birth years and months as well as age of individuals. We compute the age in months using the birth year and month of individuals and interview time. There is around 1% of individuals, who do not remember their age. We exclude these people from the analysis sample. Using information on birth years and months instead of information on age to compute age can avoid the measurement error caused by the heaping problem.

Pischke, 2009; Heckman et al., 1999). For example, education is an important determinant of employment, but we do not control for it since it can be affected by cash transfers (see the reviews by Baird et al., 2014; Millán et al., 2019; Churchill et al., 2021). For robustness analyses, we also employ a model without control variables as well as a model with a higher number of control variables.

In an RDD, choosing a suitable bandwidth around the cutoff point is an important issue. A large bandwidth can reduce the standard error of the estimates because of a higher number of observations but this choice comes with a cost in terms of misspecification bias. The RDD estimates the local effect of the treatment around the cutoff point. A small bandwidth can reduce the bias, but it also reduces estimation efficiency. In this study, we mainly rely on a bandwidth of 5 years or 60 months. It means that we limit the sample to individuals aged 15 to 64 who live in a household where the oldest member has age 75 to 84. We also examine the sensitivity of the estimates to different bandwidths and find that the different bandwidths give quite similar estimates.

A concern with the RDD in our study is that individuals living with older people aged around the cutoff point might not be representative for the whole population. In Table A.1 in the Appendix, we compare the mean of the outcome and control variables between all the people aged 15-64 and those (also aged 15-64) living with the oldest member being of age 75-84. Overall, the means of the variables of people living with the oldest member aged 75-84 are very similar to those of all the people.

We apply the same regression model (1) for different outcomes (see the outcome variables listed in Table 1). It should be noted that although several outcome variables are binary, we use a linear probability model instead of probit or logit models. Our model

contains interaction terms, which are not estimated correctly in nonlinear models. Moreover, appropriate estimators for a binary model with endogenous binary variables are not available. The 2SLS estimators are still consistent and widely used for the binary model with binary endogenous variables (Angrist 2001; Angrist and Krueger 2001).

Regarding standard errors, a common approach is to cluster standard errors at the level of the running variables that is age in months of the oldest member. In addition, we also cluster the standard error at the household level. We adopt the multiway clustering technique of Cameron et al. (2011), which allows us to cluster standard errors at both the age of the oldest and the household level.⁶ Our estimates are robust to different ways of clustering standard errors.

A main weakness of a RDD relying on age-based discontinuities is the problem of inevitable treatment with the passage of time (Lee and Lemieux, 2010). Future eligible individuals can anticipate the cash transfer receipt in the coming years, and as a result they may change their behavior before the transfer receipt. Lee and Lemieux (2010) give an example that social security payment may not have a significant effect on consumption at the age cutoff because people can smoothen their consumption according to the prediction of life-cycle theories under the assumption of no liquidity constraints. This issue can violate the no-anticipation assumption in impact evaluation which assumes that the treatment group does not change their behavior in expectation of the treatment date (Abbring and van den Berg, 2003). Although we cannot fully investigate this issue, we do not believe it represents a serious issue in our study. People in a lower middle-income country like Vietnam often face liquidity

⁶ This can be done easily using the command ivreg2 in Stata.

constraints and would not change their current labor supply because of the future transfer. In addition, as we will see in Section 4, we perform a 'donut' RDD, in which we drop households around the cutoff point, and still find similar effects of the cash transfer receipt on individuals' employment. We also perform a placebo test of the instrumental variable by using a number of different cutoff points close to 80, and do not find a significant effect at these cutoff points.

4. The impact of cash transfers on employment

In this section we present and discuss our results, beginning with the effect of the cash transfer program on individuals' employment and households' income. Next, we present a series of falsification and robustness analyses to verify our estimates.

4.1. The impact of cash transfers on employment

The main assumption on which the RDD relies is the continuity of the running variable around the cutoff point. In our study, although the actual age cannot be manipulated, it is possible that some people might report an older age to be eligible for the cash transfer. Thus, it is important to conduct the density test in the data sample. In the pooled data of the VHLSSs 2012 to 2020, the number of observations at age 78, 79, 80 and 81 is very close, at 2,079, 1,913, 1,942, and 1,928, respectively. We first conduct a simple Bernoulli test that the probability of age 79 and age 80 within this age group equals 1/2 (Cattaneo et al., 2019). The p-value of the two-side test is 0.50, which indicates that there is no evidence of "sorting" around the threshold. Next, we graph the density of age in years and age in months in Figure A.3 in the Appendix. It shows that there is no spike or jump right after the threshold of age 80 years or 960 months. More formally, we conducted a manipulation test, developed by

McCrary (2008), using the 'rddensity' Stata code (Cattaneo et al., 2018). The test statistics is equal to -0.08 (P-value of 0.93) and 1.43 (P-value of 0.15) age in years and months, respectively. Overall, we do not find statistical evidence of systematic manipulation at the age around 80.

We now tunr to the first-stage regressions of the cash transfer receipt on the instrumental variable and control variables. As mentioned, we focus on the sample of individuals aged 15-64, who live with the oldest household member aged from 75 to 84. It means that the bandwidth is 60 months around the cutoff of 960 months of the oldest member. For robustness check, we also use the optimal (data-driven) bandwidth in the nonparametric estimator and the fixed bandwidth of 36 months and 84 months in the parametric regression. Table 2 shows a strong and significant effect of the instrumental variable 'the oldest household member aged 960 months or over' on the probability of living in a household receiving the cash transfer. For the bandwidth of 60 months (column 2 in Table 2), having the oldest household member at the cutoff point increases the probability of receiving the cash transfer by 0.24. The statistics from Cragg-Donald and Kleibergen-Paap weak identification tests are very high, indicating that the instrument is very strong (see last rows in Table 3).⁷ Regarding other control variables, we find that age, gender and ethnicity of individuals are not statistically significant. However, people in urban areas are less likely to receive the cash transfers.

[Table 2 around here]

⁷ According to Staiger and Stock (1997), an F-statistic below 10 indicates weak instruments.

Table 3 presents the impact of the cash transfer receipt of households on their members' employment. Overall, the transfer receipt has no significant effect on the probability of individuals working or the number of hours they work. The point estimates are also very small. The finding that there is no disincentive effect of cash transfers on labor force participation is consistent with previous studies such as Kuhn et al. (2011), Barrientos and Villa (2015), Banerjee et al. (2017), Baird et al. (2018), Ham and Michelson (2018), Jones and Marinescu (2022), and Vera-Cossio (2022).⁸

Interestingly, although the cash transfer program has no significant effects on the labor force participation, it has a significant effect on the probability of being employed. Living in families receiving support from the cash transfer program increases the probability of having a wage-paying job by 0.14. To further examine this issue, we classify wage-paying jobs into public and private sector jobs, respectively, then use the same estimation strategy to estimate the effect of the transfer on public and private wage-paying jobs separately. The results reported in Table A.2 in the Appendix show that the cash transfer program increases the probability of having a wage-paying job in the private sector but not in the public sector. The effect of the transfer program on the probability of being employed in the public sector is of smaller magnitude and not statistically significant. This result is expected, since the opportunity of a job in the public sector is limited, especially in the short run. Similarly, we do not find a significant effect of the transfer program on having a formal job, which is a job associated with social insurance.

[Table 3 around here]

⁸ For a review of the impact of cash transfers, see Kabeer and Waddington (2015), Parker and Todd (2017), Baird et al. (2018) and Bastagli et al. (2019).

For people with a wage job, column 7 (Table 3) shows the effect of the cash transfer on their wages. The estimate is not statistically significant at the conventional level. This effect is estimated for wage workers only. If we use the full sample of people including selfemployed people and use transformation 'log of (wage + 1)', then we find a positive effect of the transfer program on per capita wage income (column 8, Table 3). This finding is consistent with the finding on the positive effect of the cash transfer on the probability of having wage-paying jobs.

The effect of the cash transfer program on self-employed farm work is very small and not statistically significant (column 5 in Table 3). However, the cash transfer program has a negative and significant effect on self-employed non-farm work. It reduces the probability of self-employed non-farm work by 0.13. It means that the cash transfer induces people to move from self-employed non-farm work to wage-paying jobs.

In Table 4, we measure the cash transfer by the amount of cash transfers instead of the dummy variable of the transfer receipt. Using the continuous variable allows us to estimate the elasticity of the employment variables with respect to the cash transfer. Although the standard RDD is applied to binary treatment variables, the RDD can be applied for continuous treatment variables in a similar context (e.g., Almond et al., 2010; Yan et al., 2020; Dong et al., 2021). Figure A.4 in the Appendix shows a discontinuity in the amount of per capita transfer income at the cutoff point in the age of the oldest member. We use the transformation log of (transfer income + 1) and apply the same model specification as in Table 3 to estimate the effect of the per capita transfer income on individuals' employment.⁹

⁹ In addition to transformation of log(x+1), we also try to transform the transfer income using the inverse hyperbolic sine transformation $(\operatorname{arsinh}(x) = \ln(x + \sqrt{1 + x^2}))$, which has a similar interpretation as the log 18

Table 3 shows similar results as the effect of the transfer receipt. The cash transfer income tends to move people from self-employed non-farm work to wage-earning jobs. A one-percent increase in the transfer income increases the probability of having a wage job by 0.02 percentage points and reduces the probability of having self-employed non-farm work by 0.02 percentage point.

[Table 4 around here]

Finally, we explore the reduced-form effect of 'the oldest household member aged 960 months (80 years) or over' on employment of household members aged 15-64 years. The reduced-form regression is reported in Table A.3 in the Appendix, showing that having a household member aged 960 months or over increases the probability of having a wage-paying job by 0.035, but reduces the probability of having a self-employed non-farm job by 0.032. We illustrate the reduced-form effect by graphing the outcome variables of individuals aged 15-64 years across the age of the oldest member in Figure 3.¹⁰ It shows a discontinuity in wage-paying job, self-employed non-farm work, and log of (wages + 1).

[Figure 3 around here]

4.2. Falsification analysis

The RDD assumes that the cutoff point can 'mimic' a randomized treatment and individuals around the cutoff point should be very similar. We therefore conduct the following falsification (placebo) tests. First, we estimate the correlation of the cash transfer program

function but can avoid a zero value of the dependent variable (e.g., see Pence (2006), and Card and DellaVigna (2020) for discussion of this transformation). The results are very similar to the transformation of log(x+1).

¹⁰ We prepare the graphs from the quadratic regressions, which are generated with the Stata command 'rdplot'. According to Gelman and Imbens (2019), local linear and local quadratic regressions should be used instead of higher-order polynomials.

with several exogenous variables (gender, age, ethnicity, education level, urban area, the number of schooling years) using the same model specification as Table 3. Table A.4 in the Appendix shows that the cash transfer program has small and insignificant estimates in all the regressions, indicating that the exogenous variables are similar around the cutoff point. In Table A.5 in the Appendix, we run reduced-form regressions of exogenous variables on the instrumental variable 'the oldest household member aged 960 months or over' on employment of household members aged 15-64 years. It shows no significant estimates in all the regressions.

Second, we employ the panel feature of the VHLSSs to examine whether individuals around the cutoff points were similar before the treatment group received the cash transfer program. The VHLSSs contain rotated panel data between two consecutive surveys. We run 2SLS regression of the baseline employment outcomes in the previous survey round on the current receipt of the cash transfer using the same model specification as Table 3. We expect that before receiving the cash transfer both the treatment and control groups are similar in the outcome variables. This expectation is confirmed in Tables A.6 and A.7 in the Appendix, which show insignificant effects of the cash transfer on the baseline outcomes.

Third, we run the reduced-form regression of the outcome variables on the instrumental variable 'having a household member aged just above 960 months' using the VHLSSs from 2004 to 2010. As mentioned, the transfer has been provided for older people aged 80 years and above since 2011. Table A.8 in the Appendix shows that the instrumental variable is not statistically significant for all the outcomes in the data before 2011. On the other hand, the instrumental variable is statistically significant at the 1% level in a regression of wage and non-farm employment in the data from 2012. These findings suggest that the

significant effect of the instrumental variables on wage and non-farm employment occurs through the cash transfer program.

Fourth, we perform a placebo test of the instrumental variable by using a number of different cutoff points for the oldest household member at 70, 75, 80 and 85 instead of 80 years old. We run the reduced-form regression of wage-paying job and self-employed non-farm work of individuals aged 15-64 on these instruments using the same mode specification as Table A.3 in the Appendix. We graph the point estimates and their 95% confidence intervals of the local effect of these age thresholds on the probability of having a wage job and the probability of non-farm work in Panels A and B of Figure 4. The sample includes individuals aged 15-64 living in households with the oldest member in different age bandwidths. The estimates are from the reduced-form regressions of employment outcomes using similar specification as Table A.3 in the Appendix. It shows that only the instrument 'the oldest household member aged 960 months or over' is statistically significant, while the effect of the instruments at other age thresholds is not significant at the 10% level.

In addition to the above thresholds, we try to use thresholds close to 80 years old, i.e., 75 to 79, to examine the 'no-anticipation assumption'. If there is an anticipation effect of the cash transfer, we would expect an increasing or decreasing effect of the instrumental variable across these age thresholds. Panels A and B of Figure 4 do not show this trend. The point estimates have a small magnitude, and all the estimates are not significant at the 10% level.

[Figure 4 around here]

4.3. Robustness analysis

We conduct several analyses to examine the robustness of the effect estimates of the cash transfer program to different model specifications and bandwidths. First, we examine

whether the effect estimate of the cash transfer receipt is sensitive to control variables. Table A.9 in the Appendix reports the 2SLS regression without covariates, and Table A.10 in the Appendix presents the 2SLS regression with a large number of covariates. Additional covariates in Table A.11 include dummies of education levels of household head, household size, the proportion of children, and the proportion of older people from 65 years and above in households, the poverty status of household, and province fixed effects.¹¹ All estimates from the small and large models are almost unchanged compared to those from the main specification model in Table 3.

Second, we use the data sample which excludes the 2012 VHLSS and the 2020 VHLSS. The 2012 year is the year before the official issue of the Decree 136 on the cash transfer program, while the 2020 year is the COVID-19 affected year. Tables A.11 and A.12 in the Appendix show similar estimates of the effect of the cash transfer on employment using the sample without the 2012 and 2020 VHLSSs.

Third, we check the robustness of the results to the weighting schemes and standard error clustering. In Table A.13 in the Appendix, we estimate the effect of the cash transfer without using the sampling weight. In Table A.14, we use different ways of clustering standard errors (one-way clustering at the age of the oldest member; one-way clustering at the village level; and two-way clustering at the age of the oldest member and the village level). Overall, the results are very similar to the main results used for interpretation (reported in Table 3).

¹¹ At the first-level administrative division, Vietnam includes 63 provinces and cities.

Fourth, related to the standard error issue, there can be a multiple testing problem in the case of several dependent variables. Traditional estimation reports the p-value of the estimate, which is the false positive rate among all the results. Alternatively, we estimate the q-values (the false discovery rate), which is the false positive rate among significant results. Figure A.5 in the Appendix graphs the p-values and q-values (estimated by the method of Simes (1986)) of the effect of the transfer receipt.¹² It shows that the effect estimates on wage jobs, self-employed non-farm work, and log of wages of all people remain significant at the 5% level.

Fifth, we try nonparametric estimation controlling for covariates developed by Calonico et al. (2014) and Calonico et al. (2019). The 'rdrobust' command from Calonico et al. (2017) allows us to estimate the local effect of the cash transfer receipt on employment outcomes using local polynomial regression. For comparison, the control variables and clustering are kept the same as those in Table 3 (the estimates without control variables are very similar). The selection of optimal bandwidths is based on a data-driven mean-squared error (MSE). For robustness, we report estimates using one common MSE-optimal bandwidth selector as well as one common MSE-optimal bandwidth selector for the sum of regression (for details see Calonico et al., 2017). For each bandwidth selection algorithm, estimates from three estimators (conventional RDD estimates; bias-corrected RDD estimates with a robust variance estimator; bias-corrected RDD estimates with a robust variance estimator) are reported. Reassuringly, the nonparametric results, which are presented in Table A.15 in the Appendix, are similar to those from the parametric approach. The point estimates

¹² These q-values are estimated using the 'qqvalue' command in Stata (see Newson, 2011).

from the nonparametric tends to be larger than those from the parametric regression. However, this is because the bandwidth selected in the parametric approach is smaller than the bandwidth used in the parametric regression. As we will see in the following paragraph, if we use similar bandwidths in the parametric regression, the estimates are more similar to those from the nonparametric approach.

Last but not least, we examine the sensitivity of the effect estimates from the parametric regression to different bandwidths. We first apply the same optimal bandwidths used in the nonparametric approach in Table A.15. The bandwidth varies across outcomes. Tables A.16 and A.17 in the Appendix show very similar estimates as the nonparametric approach. Next, we use the fixed bandwidths of 36 months (3 years) and 84 months (7 years) for estimation. The results, reported in A.18 and A.19 in the Appendix, show similar findings that people in households receiving the cash transfer tend to move from self-employed nonfarm work to wage-paying jobs. We vary the bandwidth from 12 to 84 months (with a 12month change) and graph the estimates of the cash transfer on the probability of wage-paying job and the probability of self-employed non-farm work across the bandwidths in Figure A.6 in the Appendix. The point estimates are all positive but tend to decrease as the bandwidth increases. As expected, the standard error is larger for smaller bandwidths. Finally, Table A.20 in the Appendix shows the 'donut' RDD, in which we drop individuals with the oldest members being very close to the cutoff point (960 months) to examine whether the effects are sensitive to the sample around this point (Cattaneo et al., 2019). The results remain almost the same as those shown in Table 3.

5. Mechanisms

Households in low- and middle-income countries are more likely to face credit and liquidity constraints than their developed country counterparts, and they can use cash transfers for their farm and non-farm production. Several studies show a positive effect of cash transfers on self-employment, especially non-farm employment, due to increasing investment from cash transfers in non-farm production and business (e.g., Lindh et al., 1996; Blattman et al., 2014; and Salehi-Isfahani and Mostafavi-Dehzooei, 2018; Vera-Cossio, 2022). No positive effect of the cash transfer program on self-employment emerges among our results. However, we find that the cash transfer program tends to help individuals' mobility from self-employed non-farm work to wage-paying jobs.

On several grounds, we argue that one of the main mechanisms through which the cash transfer program moves adult individuals from self-employed non-farm work to wage-paying jobs can be increased childcare provided by older household members. First, we show that the cash transfer program reduces the labor force participation of older people. So far, we have focused on the effect of the transfer program on employment of individuals aged 15-64. Table 5 presents a 2SLS regression of labor supply of people aged from 65 on the cash transfer receipt. We do not report the regression of wage-paying jobs and formal jobs, since the effect is not statistically significant and there was only 3.5% and 0.2% of the older people having wage-paying jobs and formal jobs, respectively. Table 5 shows that receipt of a cash transfer reduces the probability of working for people aged from 65 years by 0.88 and those from 75 years by 0.11 (Table A.21 in the Appendix presents the regression using other bandwidths). The effect estimates are similar for the probability of having farm self-employment, suggesting that the transfer program reduces the labor force participation of

older people with farm self-employment. Earnings from farm work are low for older people, and the cash transfer can compensate for their earnings loss when not working.

[Table 5 around here]

Second, we argue that when older people quit working, they can take care of small children for younger adults to find wage-paying jobs. People can work at home and take care of children at the same time, but if they want to have a wage job, they need someone to take care of their children. In Vietnam, it is quite common that grandparents live with children and take care of their grandchildren. It is estimated from the 2020 VHLSS that 30% of children below 6 years old live with grandparents. A number of recent studies show a positive effect on women's labor supply of formal childcare (e.g., Bauernschuster and Schlotter, 2015; Martínez and Perticará, 2017; Dang et al., 2022) as well as informal childcare from grandparents (e.g., Maurer-Fazio et al., 2011; Zanella, 2017; Garcia-Moran, E., & Kuehn, 2017; Halim et al., 2021). Recently, Rupert and Zanella (2018) and Frimmel et al. (2022) show that becoming grandparents strongly decreases the labor force participation of older people. For the case of Vietnam, Dang et al. (2022) find a large and positive effect of formal childcare on the probability of wage-paying jobs of mothers. There are no studies on the effect of grandparents' childcare on adults' employment. To examine this issue, we run an OLS regression of variables of wage-paying job and self-employed non-farm work on the proportion of members aged 65 years and above in households. Table A.22 in the Appendix shows that the proportion of older members has a positive correlation with the probability of having a wage job and a negative correlation with the probability of having self-employed non-farm work in households with children below 6 (column 3). In households with small children, the presence of older people is strongly associated with the probability of having a wage job of younger females (column 7).

We run a 2SLS regression of the wage job on the cash transfer receipt separately for different population subgroups using the same model specification as in Table 3. Figure 5 presents the point estimate and the 95% confidence interval of the estimated effect of the transfer program on the probability of having a wage-paying job across different population subgroups. Of particular interest is the effect on individuals with children and women. There is a stronger effect of the transfer program on people living with children than those not living with children. The effect of the transfer program is statistically significant for women in the fertility age but not for men. These findings are consistent with the argument that childcare helps adults, especially women to increase their labor market participation.

The transfer program has a larger effect on rural and ethnic minority people than urban and Kinh people. Possibly, formal childcare (children from kindergarden and childcare centers) is limited in the rural and ethnic minority areas, and as a result, informal childcare by older people is more important for these groups. Older people are more likely to reduce the labor supply to provide childcare for adults to find a wage-paying job.

[Figure 5 around here]

The effect of the cash transfer on the probability of having a wage job is not statistically significant for people with a college or university degree. Most of these people already have a wage job. The effect of the cash transfer is highest for those with uppersecondary or vocational degree. This finding is consistent with the finding on the positive effect of the cash transfer on labor mobility from non-farm self-employment to wage jobs. In our study, people with wage jobs have the highest education level, while those with farm self-employment have the lowest education level. Thus, it might be more difficult for selfemployed people in the farm sector to move to wage-paying jobs than for people with nonfarm self-employment.

Education including preschool can be also affected by cash transfers (see reviews by Baird et al., 2014; Millán et al., 2019; Churchill et al., 2021). If cash transfers are used for formal childcare of children, parents can also increase their labor market participation. Table A.23 in the Appendix reports a 2SLS regression of school enrolment on the cash transfer program for children of different age groups. The coefficient of the cash transfer is positive in all the regressions but not statistically significant at conventional levels. Thus, there is no evidence that the cash transfer increases the labor market participation of adults through improving formal healthcare for children.

6. Conclusions

Evidence on the impact and mechanism of cash transfers on labor supply is very useful for designing effective cash transfer programs. This study has examined the effect of cash transfers from the social assistance policy on individuals' employment in Vietnam using a RDD and recent household data. We find mixed evidence on the impact of the cash transfer on labor supply. There are no significant effects of households' receipt of a cash transfer on labor force participation as well as working hours of individuals aged 15-64. However, there is a disincentive effect on the labor force participation of older people. Importantly, we find robust evidence that cash transfer receipts increase younger people's mobility from self-employed non-farm work to wage-paying jobs. We argue that this effect happens because the

cash transfer reduces the labor force participation of older people, who can provide childcare helping younger parents to find wage-paying jobs.

There are several policy implications from our study. First, provision of cash transfers for vulnerable groups is an important policy. For the case of Vietnam, there is no evidence of a disincentive effect of the cash transfer program on individuals aged 15-64, who are the main labor force. The cash transfer can compensate for earning loss of older people if they want to reduce the labor supply. Second, although the cash transfer can be received mainly by older people, it can change the labor behavior of younger people. This finding reflects a spill-over effect of cash transfers. Third, our study provides suggestive evidence on the important role of childcare in general and informal childcare from older people in particular in helping younger parents to participate in the labor market and find wage-paying jobs.

References

Abel, M. (2019). Unintended labor supply effects of cash transfer programs: New evidence from South Africa's pension. *Journal of African Economies*, 28(5), 558-581.

Ai, C., & Norton, E. C. (2003). Interaction terms in logit and probit models. *Economics Letters*, 80(1), 123–129.

Almond, D., Doyle, J. J., Kowalski, A. E., and Williams, H. (2010), Estimating marginal returns to medical care: Evidence from at- risk newborns. *The Quarterly Journal of Economics*, 125, 591–634.

Angrist, J. D. (2001). Estimation of limited dependent variablemodels with dummy endogenous regressors: Simple strategies for empirical practice. *Journal of Business and Economic Statistics* 19(1), 2–28.

Angrist, J. D., & Pischke, J. (2009). *Mostly Harmless Econometrics: An Empiricist's Companion*. New Jersey: Princeton University Press.

Angrist, J. D., and A. B. Krueger. (2001). Instrumental variables and the search for identification: From supply and demand to natural experiments. *Journal of Economic Perspectives* 15(4), 69–85.

Atalay, K., & Barrett, G. F. (2014). The causal effect of retirement on health: New evidence from Australian pension reform. *Economics Letters*, *125*(3), 392-395.

Baird, S., Ferreira, F. H., Özler, B., & Woolcock, M. (2014). Conditional, unconditional and everything in between: a systematic review of the effects of cash transfer programs on schooling outcomes. *Journal of Development Effectiveness*, 6(1), 1-43.

Baird, S., McKenzie, D., & Özler, B. (2018). The effects of cash transfers on adult labor market outcomes. *IZA Journal of Development and Migration*, 8(1), 1-20.

Banerjee, A. V., Hanna, R., Kreindler, G. E., & Olken, B. A. (2017). Debunking the stereotype of the lazy welfare recipient: Evidence from cash transfer programs. *The World Bank Research Observer*, *32*(2), 155-184.

Barrientos, A., & Villa, J. M. (2015). Antipoverty transfers and labour market outcomes: Regression discontinuity design findings. *The Journal of Development Studies*, *51*(9), 1224-1240.

Bastagli, F., Hagen-Zanker, J., Harman, L., Barca, V., Sturge, G., & Schmidt, T. (2019). The impact of cash transfers: a review of the evidence from low-and middle-income countries. *Journal of Social Policy*, *48*(3), 569-594.

Bauernschuster, S., & Schlotter, M. (2015). Public childcare and women' labor supply— Evidence from two quasi-experiments. *Journal of Public Economics*, 123, 1–16.

Blattman C., Fiala N., Martinez S. (2014). Generating skilled self-employment in developing countries: Experimental evidence from Uganda. *Quarterly Journal of Economics*, 129 (2): 697–752.

Calonico, S., Cattaneo Matias D., Farrell Max H., and Titiunik Rocio. (2018). Regression discontinuity designs using covariates. *The Review of Economics and Statistics* 101(3): 442–451.

Calonico, S., Cattaneo, M. D., Farrell, M. H., & Titiunik, R. (2019). Regression discontinuity designs using covariates. *Review of Economics and Statistics*, *101*(3), 442-451.

Calonico, S., Cattaneo, M.D., Farrell, M.H., Titiunik, R., (2017). Rdrobust: software for regression discontinuity designs. STATA J. 17 (2), 372–404.

Calonico, S., Matias D. Cattaneo, and Rocio Titiunik, (2014). Robust nonparametric confidence intervals for regression-discontinuity designs, *Econonetrica*, 2014, 82 (6), 2295–2326.

Cameron, A. Colin, Jonah Gelbach, and Douglas Miller. (2011). Robust inference with multiway clustering. *Journal of Business and Economic Statistics* 29 (2): 238–49.

Card, D., and S. DellaVigna. (2020). What do editors maximize? Evidence from four economics journals. *The Review of Economics and Statistics* 102 (1): 195–217.

Card, David, David S. Lee, Zhuan Pei, and Andrea Weber, (2017). Regression kink design: Theory and practice, in Matias D. Cattaneo and Juan Carlos Esanciano, eds., Advandse in Econometrics, Vol. 38, Oxford University Press, 2017.

Cattaneo, M. D., Idrobo, N., & Titiunik, R. (2019). A practical introduction to regression discontinuity designs: Foundations. Cambridge University Press.

Cattaneo, M. D., M. Jansson, and X. Ma. 2018. Manipulation testing based on density discontinuity. *Stata Journal* 18, no. 1:234–61.

Churchill, S. A., Iqbal, N., Nawaz, S., & Yew, S. L. (2021). Unconditional cash transfers, child labour and education: theory and evidence. *Journal of Economic Behavior & Organization*, 186, 437-457.

Dang, H. A. H., Hiraga, M., & Nguyen, C. V. (2022). Childcare and maternal employment: evidence from Vietnam. *World Development*, *159*, 106022.

Dang, H. A. H., Nguyen, C. V., & Carletto, C. (2023). Did a Successful Fight against COVID-19 Come at a Cost? Impacts of the pandemic on employment outcomes in Vietnam. *World Development*, 161, 106129.

Dong, Y., Lee, Y. Y., & Gou, M. (2021). Regression discontinuity designs with a continuous treatment. *Journal of the American Statistical Association*, 1-14.

Fan, M., He, G., & Zhou, M. (2020). The winter choke: coal-fired heating, air pollution, and mortality in China. *Journal of Health Economics*, 71, 102316.

Farrington, J. and Slater, R. (2006). Introduction: Cash transfers: Panacea for poverty reduction or money down the drain?, *Development Policy Review* 24 (5): 499-511.

Frimmel, W., Halla, M., Schmidpeter, B., & Winter-Ebmer, R. (2022). Grandmothers' labor supply. *Journal of Human Resources*, *57*(5), 1645-1689.

Garcia-Moran, E., & Kuehn, Z. (2017). With strings attached: Grandparent-provided child care and female labor market outcomes. *Review of Economic Dynamics*, 23, 80-98.

Gelman, A., & Imbens, G. (2019). Why high-order polynomials should not be used in regression discontinuity designs. *Journal of Business & Economic Statistics*, 37(3), 447–456.

Gertler, P. J., Martinez, S. W., & Rubio-Codina, M. (2012). Investing cash transfers to raise long-term living standards. *American Economic Journal: Applied Economics*, 4(1), 164-92.

Gorry, A., Gorry, D., & Slavov, S. N. (2018). Does retirement improve health and life satisfaction?. *Health economics*, 27(12), 2067-2086.

Government of Vietnam (2000), Nghị định Quy định chính sách cứu trợ giúp xã hội, Nghị định số 07/2000/NĐ-CP ban hành ngày ngày 09 tháng 03 năm 2000 (In English: Decree 07/2000/NĐ-CP of the Government of Vietnam on social support policies, issued on 09/03/2000).

Government of Vietnam (2007), Nghị định Quy định chính sách trợ giúp xã hội đối với đối tượng bảo trợ xã hội, Nghị định số 67/2007/NĐ-CP ban hành ngày ngày 13 tháng 04 năm 2007 (In English: Decree 67/2007/NĐ-CP of the Government of Vietnam on social support policies for social protection beneficiaries issued on 13/04/2007).

Government of Vietnam (2011), Nghị định Quy định chi tiết và hướng dẫn thi hành một số điều của Luật Người cao tuổi, Nghị định số 06/2011/NĐ-CP ban hành ngày ngày 14 tháng 01 năm 2011 (In English: Decree 06/2011/NĐ-CP of the Government of Vietnam on detailed regulations of the Law on the Elderly, issued on 14/01/2011).

Government of Vietnam (2013), Nghị định Quy định chính sách trợ giúp xã hội đối với đối tượng bảo trợ xã hội, Nghị định số 136/2013/NĐ-CP ban hành ngày ngày 21 tháng 10 năm

2013 (In English: Decree 136/2013/NĐ-CP of the Government of Vietnam on social support policies for social protection beneficiaries issued on 21/10/2013).

Hahn, J., Todd, P., & Van der Klaauw, W. (2001). Identification and estimation of treatment effects with a regression-discontinuity design. *Econometrica*, 69(1), 201–209.

Halim, D., Perova, E., & Reynolds, S. (2021). Childcare and mothers' labor market outcomes in lower and middle-income countries, Policy Research Working Paper No. 9828. World Bank, Washington, DC.

Ham, A., & Michelson, H. C. (2018). Does the form of delivering incentives in conditional cash transfers matter over a decade later?. *Journal of Development Economics*, *134*, 96-108.

Handa, S., Daidone, S., Peterman, A., Davis, B., Pereira, A., Palermo, T., & Yablonski, J. (2018). Myth-busting? Confronting six common perceptions about unconditional cash transfers as a poverty reduction strategy in Africa. *The World Bank Research Observer*, 33(2), 259-298.

Heckman, J., LaLonde, R., & J; and Smith Jeffrey A. (1999). The economics and econometrics of active labor market programs. In A. Ashenfelter & D. Card (Eds.). Handbook of Labor Economics 1999 (volume 3). Elsevier Science.

Imbens, G. W., & Lemieux, T. (2008). Regression discontinuity designs: A guide to practice. *Journal of Econometrics*, 142(2), 615-635.

Jones, D., and Marinescu, I. (2022). The labor market impacts of universal and permanent cash transfers: Evidence from the Alaska Permanent Fund. *American Economic Journal: Economic Policy*, 14 (2): 315-40.

Kabeer, N., and Waddington, H. (2015). Economic impacts of conditional cash transfer programmes: A systematic review and meta-analysis. *Journal of Development Effectiveness*, 7(3), 290-303.

Kuhn, P., Kooreman, P., Soetevent, A.R. and Kapteyn, A. (2011). The own and social effects of an unexpected income shock: Evidence from the Dutch Postcode Lottery. *American Economic Review*, vol. 101, pp. 2226–2247.

Lee, D. S., & Lemieuxa, T. (2010). Regression discontinuity designs in economics. *Journal of Economic Literature*, 48(2), 281–355.

Lindh, T., and Ohlsson, H. (1996). Self-Employment and windfall gains: Evidence from the Swedish Lottery. *Economic Journal* 106 (439): 1515–26.

Lloyd-Sherlock, P. (2006). Simple transfers, complex outcomes: The impacts of pensions on poor households in Brazil. *Development and Change* 37 (5): 969-95.

Martínez, C., & Perticará, M. (2017). Childcare effects on maternal employment: Evidence from Chile. *Journal of Development Economics*, 126, 127–137.

Maurer-Fazio, M., Connelly, R., Chen, L., & Tang, L. (2011). Childcare, eldercare, and labor force participation of married women in urban China, 1982–2000. *Journal of Human Resources*, 46(2), 261–294.

McCrary, J. (2008). Manipulation of the running variable in the regression discontinuity design: A density test. *Journal of Econometrics*, 142(2), 698–714.

Millán, T. M., Barham, T., Macours, K., Maluccio, J. A., & Stampini, M. (2019). Long-term impacts of conditional cash transfers: review of the evidence. *The World Bank Research Observer*, *34*(1), 119-159.

MOLISA (2022), Summary of 10 years of implementation of the Resolution of the 5th Plenum of the XI Central Committee related to some issues of social policy in the period 2012 – 2020, Ministry of Labor - Invalids and Social Affairs, Vietnam. Available at: http://www.molisa.gov.vn/Pages/tintuc/chitiet.aspx?tintucID=232799

Newson, R. B. (2011). Frequentist Q-values for multiple-test procedures. *The Stata Journal*, 10(4), 568–584.

Nguyen, C. (2013). The impact of social security on household welfare: Evidence from a transition country. *The European Journal of Development Research*, 25(5), 737-757.

Nguyen, C. (2021). Old-Age Pensions in a Lower Middle-Income Country: Economic or Psychological Effects?. *Economic Development and Cultural Change*, 69(3), 1165-1202.

Pence, K. M. (2006). The Role of Wealth Transformations: An Application to Estimating the Effect of Tax Incentives on Saving. *Contributions in Economic Analysis & Policy* 5 (1). doi:10.2202/1538-0645.1430.

Rawlings, L.B., and Rubio, G.M. (2005). Evaluating the impact of conditional cash transfer programs. *The World Bank Research Observer*, 20(1), 29-55.

Rose, L. (2020). Retirement and health: Evidence from England. *Journal of Health Economics*, 73, 102352.

Rupert, P., & Zanella, G. (2018). Grandchildren and their grandparents' labor supply. *Journal of Public Economics*, 159, 89-103.

Salehi-Isfahani, D., & Mostafavi-Dehzooei, M. H. (2018). Cash transfers and labor supply: Evidence from a large-scale program in Iran. *Journal of Development Economics*, *135*, 349–367.

Simes, R. J. (1986). An improved bonferroni procedure for multiple tests of significance. *Biometrika*, 73, 751–754.

Staiger, D., and J. Stock. (1997). Instrumental variables regression with weak instruments. *Econometrica* 65(3), 557–86.

Van de Walle, D. (2004). Testing Vietnam's public safety net. *Journal of Comparative Economics* 32 (4): 661-79.

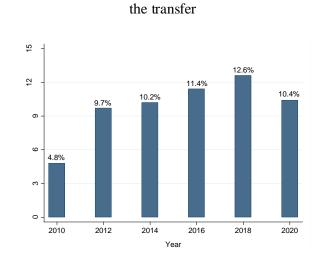
Van den Berg, M., and Nguyen, C. (2011). Impact of public and private cash transfers on poverty and inequality: evidence from Vietnam. *Development Policy Review*, 29(6), 689-728.

Vera-Cossio, D. A. (2022). Dependence or constraints? Cash transfers and labor supply. *Economic Development and Cultural Change*, 70(4), 1439-1477.

Zanella, G. (2017). How does grandparent childcare affect labor supply? *IZA World of Labor*. DOI 10.15185/izawol.337.

Zhu, R. (2016). Retirement and its consequences for women's health in Australia. *Social Science & Medicine*, *163*, 117-125.

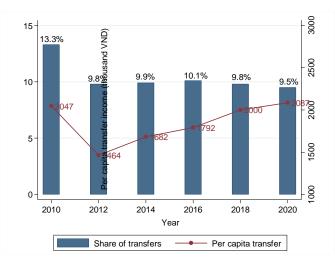
Figures and Tables



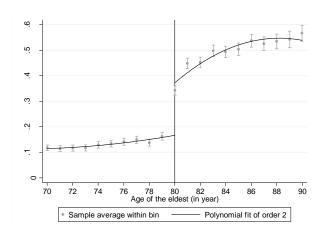
Panel A. The proportion of households receiving

Figure 1: Cash transfers for social assistance beneficiaries

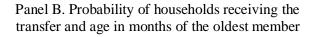
Panel B. Per capita transfer income and the share of the transfer in household income

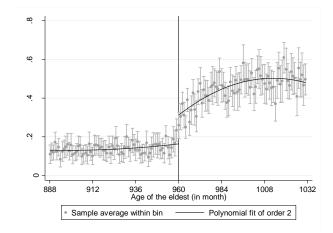


Source: Estimation from VHLSSs 2010 to 2020.



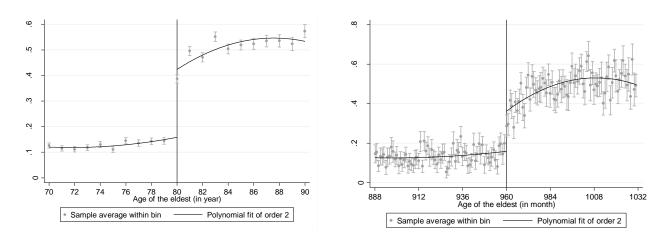
Panel A. Probability of households receiving the transfer and age in years of the oldest member





Panel A. Probability of individuals (aged 15-64) living in transfer-receiving households and age in years of the oldest member

Panel B. Probability of individuals (aged 15-64) living in transfer-receiving households and age in months of the oldest member



Note: Panels A and B graph the regression discontinuity plot of the probability of households receiving the cash transfers across age (in year and in month) of the oldest household member. The graph presents binned sample means at each age (in year and in month) and the 95% confidence interval. Panels A and B use household-level data and limit to the sample of households with the oldest household member aged 70 to 90 years.

Panels C and D show the regression discontinuity plot of the probability of individuals living in a household with the transfer across age (in year and in month) of the oldest household member. In these panels, the sample is limited to individuals (aged 15-64) living in a household with the oldest household member aged 70 to 90 years.

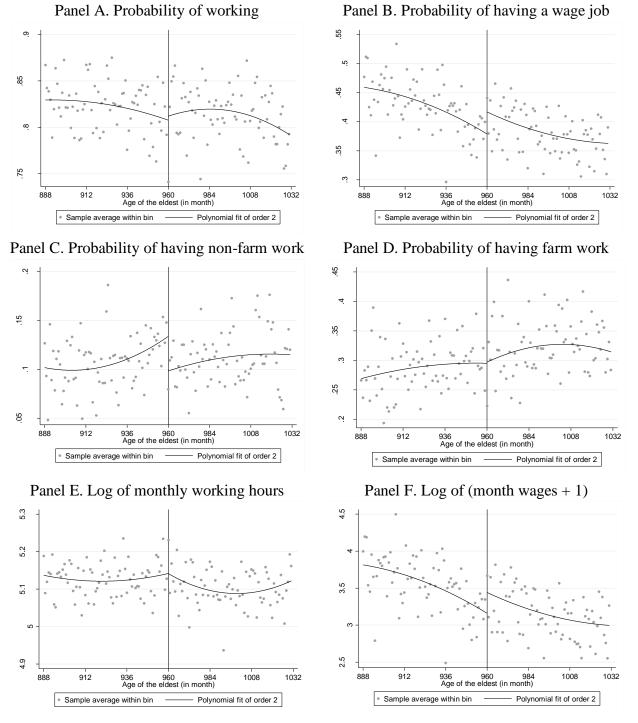
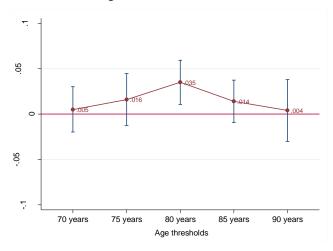


Figure 3. RDD graphs of employment outcomes

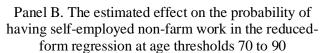
Note: This graph shows the regression discontinuity plot of employment outcomes of individuals across age of the oldest household member.

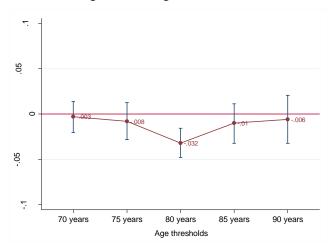
The sample is limited to individuals (aged 15-64) living in a household with the oldest household member aged 74 to 85 years.

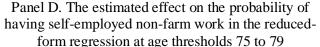
Panel A. The estimated effect on the probability of having a wage job in the reduced-form regression at age thresholds 70 to 90

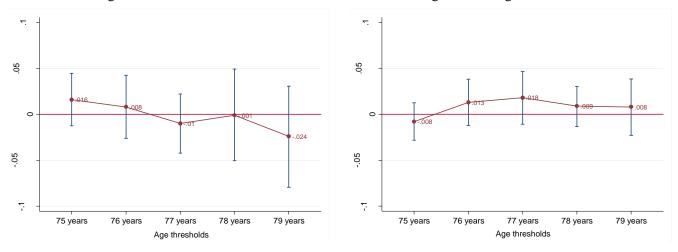


Panel C. The estimated effect on the probability of having a wage job in the reduced-form regression at age thresholds 75 to 79









Note: The figure graphs the point estimates and their 95% confidence intervals of the local effect of the age thresholds on the probability of having a wage job and the probability of non-farm work. The sample includes individuals aged 15-64 living in households with the oldest member in different age bandwidths. The estimates are from the reduced-form regressions of employment outcomes using similar specification as Table A.3 in the Appendix.

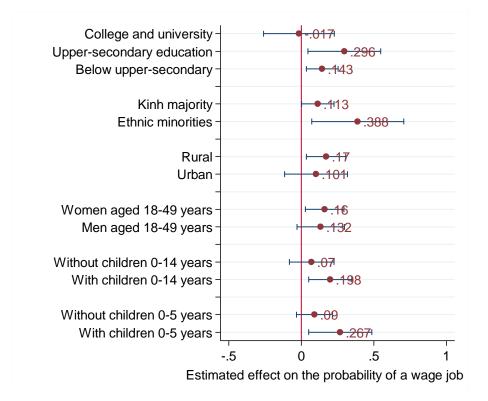


Figure 5: Heterogenous effect of the transfer on the probability of having a wage job

Note: This figure graphs the estimates and their 95% confidence intervals of the receipt of the transfer in regressions of having a wage job in different population sub-groups. The model specification is the same as Table 3.

| Outcomes | All | people aged 1 | 5-64 | People aged | l 15-64 in hous the transfer | seholds with | People aged | 15-64 in hous the transfer | seholds with |
|------------------------------------|-----------|---------------|-----------|-------------|---------------------------------|--------------|-------------|-------------------------------|--------------|
| | Year 2012 | Year 2016 | Year 2020 | Year 2012 | Year 2016 | Year 2020 | Year 2012 | Year 2016 | Year 2020 |
| Proportion of working (%) | 81.2 | 82.1 | 81.9 | 81.1 | 82.1 | 82.4 | 82.4 | 81.8 | 77.4 |
| | (0.2) | (0.2) | (0.2) | (0.2) | (0.2) | (0.2) | (0.6) | (0.6) | (0.5) |
| Proportion of having a wage job | 37.8 | 40.4 | 47.4 | 38.0 | 40.9 | 48.0 | 35.3 | 35.4 | 41.2 |
| (%) | (0.3) | (0.3) | (0.3) | (0.3) | (0.3) | (0.4) | (0.9) | (1.0) | (0.7) |
| Proportion of having self-employed | 11.2 | 13.2 | 15.0 | 11.6 | 13.8 | 15.4 | 6.4 | 7.9 | 11.6 |
| non-farm work (%) | (0.2) | (0.2) | (0.3) | (0.2) | (0.2) | (0.3) | (0.4) | (0.4) | (0.5) |
| Proportion of having self-employed | 32.2 | 28.4 | 19.5 | 31.5 | 27.3 | 19.0 | 40.7 | 38.4 | 24.5 |
| farm work (%) | (0.5) | (0.5) | (0.4) | (0.5) | (0.5) | (0.4) | (1.2) | (1.4) | (0.8) |
| Proportion of having a job with | 13.7 | 16.1 | 22.4 | 14.5 | 16.9 | 23.5 | 5.7 | 8.4 | 12.3 |
| social insurance (%) | (0.3) | (0.3) | (0.4) | (0.3) | (0.3) | (0.4) | (0.3) | (0.5) | (0.5) |
| The number of monthly working | 183.4 | 185.7 | 193.0 | 184.5 | 186.7 | 193.9 | 171.4 | 176.7 | 183.2 |
| hours | (0.7) | (0.6) | (0.6) | (0.7) | (0.6) | (0.6) | (1.7) | (1.9) | (1.4) |
| Monthly wage (thousand VND, | 4208.3 | 5212.0 | 6634.1 | 4317.6 | 5315.4 | 6760.9 | 2927.8 | 4131.6 | 5158.3 |
| computed for wage workers) | (49.4) | (52.8) | (65.7) | (51.4) | (54.6) | (67.7) | (65.8) | (88.5) | (113.7) |
| Per capita monthly wage (thousand | 1590.6 | 2105.4 | 3143.4 | 1642.0 | 2176.4 | 3245.0 | 1032.7 | 1463.9 | 2127.7 |
| VND, computed for all people) | (26.0) | (30.4) | (44.5) | (27.5) | (31.9) | (46.9) | (31.7) | (48.1) | (55.3) |

Table 1. Employment outcomes

Note: Monthly wages are adjusted to the January 2020 price using annual CPI. Poor households are those who are identified by local authorities as the poor.

Standard errors of means in parentheses. Source: Estimation from VHLSSs 2012, 2016 and 2020.

| | | e is the dummy indicated in the transformed structure of the transforme | |
|-------------------------------------|----------------------------------------------|-------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------------|----------------------------------------------|
| Explanatory variables | The sample with bandwidth of 36 months | The sample with bandwidth of 60 months | The sample with bandwidth of 84 months |
| | (1) | (2) | (3) |
| I{Oldest's month age>=960} | 0.2094*** | 0.2426*** | 0.2656*** |
| | (0.0255) | (0.0208) | (0.0183) |
| (Oldest's month age-960)*I{Eldest's | 0.0036*** | 0.0024*** | 0.0013*** |
| month age>=960} | (0.0011) | (0.0006) | (0.0004) |
| (Oldest's month age-960) | 0.0010 | 0.0004 | 0.0004** |
| | (0.0007) | (0.0003) | (0.0002) |
| Age | -0.0011 | -0.0007 | -0.0003 |
| | (0.0016) | (0.0012) | (0.0009) |
| Age squared | 0.0000 | 0.0000 | 0.0000 |
| | (0.0000) | (0.0000) | (0.0000) |
| Male (male=1, female=0) | -0.0023 | 0.0002 | -0.0008 |
| | (0.0040) | (0.0035) | (0.0029) |
| Ethnic minorities (yes=1, no=0) | 0.0032 | 0.0006 | -0.0025 |
| | (0.0202) | (0.0151) | (0.0130) |
| Urban areas (urban=1, rural=0) | -0.0955*** | -0.0980*** | -0.0954*** |
| | (0.0129) | (0.0100) | (0.0080) |
| Year fixed effects | Yes | Yes | Yes |
| Constant | 0.1058*** | 0.1020*** | 0.1052*** |
| | (0.0316) | (0.0228) | (0.0197) |
| Observations | 17,867 | 29,401 | 40,245 |
| R-squared | 0.153 | 0.175 | 0.185 |

Table 2: First-stage regression of receiving the transfer

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels. *** p < 0.01, ** p < 0.05, * p < 0.1. Source: Estimation from VHLSSs from 2012 to 2020.

| | | | | Depende | ent variables | | | |
|--------------------------------------------------------------------|------------------------------------------|-------------------------------------|--------------------------------------------------------|----------------------------------------------------|-------------------------------------------------------|----------------------------------------------------------------------------------|---------------------------------------------|--------------------------------------------|
| Explanatory variables | Currently working (yes=1, no=0) | Having wage job (yes=1, no=0) | Self- employed non-farm work (yes=1, no=0) | Self- employed farm work (yes=1, no=0) | Having social insurance (yes = 1, no = 0) | Log of number of working hours in the last month (working people) | Log of monthly wage (wage workers) | Log of monthly wage (all workers) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | 0.0158 | 0.1439*** | -0.1308*** | 0.0026 | -0.0119 | -0.0224 | -0.1094 | 1.1113** |
| | (0.0367) | (0.0529) | (0.0368) | (0.0487) | (0.0482) | (0.0804) | (0.1683) | (0.4501) |
| (Oldest's month age-960)*I{Oldest's | 0.0002 | -0.0001 | 0.0004 | -0.0001 | 0.0003 | -0.0004 | -0.0008 | -0.0008 |
| month age>=960} | (0.0003) | (0.0004) | (0.0003) | (0.0004) | (0.0004) | (0.0006) | (0.0014) | (0.0034) |
| (Oldest's month age-960) | -0.0001 | -0.0010*** | 0.0005*** | 0.0005** | -0.0003 | 0.0004 | 0.0007 | -0.0084*** |
| | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0003) | (0.0004) | (0.0009) | (0.0021) |
| Age | 0.0831*** | 0.0605*** | 0.0128*** | 0.0097*** | 0.0340*** | 0.0570*** | 0.0361*** | 0.5153*** |
| | (0.0013) | (0.0014) | (0.0010) | (0.0012) | (0.0010) | (0.0033) | (0.0056) | (0.0117) |
| Age squared | -0.0010*** | -0.0008*** | -0.0001*** | -0.0000*** | -0.0005*** | -0.0008*** | -0.0005*** | -0.0070*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0002) |
| Male (male=1, female=0) | 0.0521*** | 0.1424*** | -0.0216*** | -0.0687*** | -0.0187*** | 0.0961*** | 0.1732*** | 1.2425*** |
| | (0.0041) | (0.0060) | (0.0035) | (0.0046) | (0.0048) | (0.0070) | (0.0158) | (0.0512) |
| Ethnic minorities (yes=1, no=0) | 0.1410*** | -0.0976*** | -0.0791*** | 0.3177*** | -0.0996*** | -0.0126 | -0.6148*** | -0.9952*** |
| | (0.0052) | (0.0082) | (0.0042) | (0.0093) | (0.0054) | (0.0130) | (0.0502) | (0.0639) |
| Urban areas (urban=1, rural=0) | -0.0725*** | 0.1021*** | 0.0424*** | -0.2169*** | 0.1437*** | 0.1391*** | 0.3893*** | 1.0235*** |
| | (0.0071) | (0.0091) | (0.0078) | (0.0073) | (0.0093) | (0.0124) | (0.0264) | (0.0779) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7638*** | -0.7167*** | -0.1334*** | 0.0863*** | -0.4134*** | 4.0497*** | 7.3884*** | -6.2649*** |
| | (0.0245) | (0.0267) | (0.0181) | (0.0228) | (0.0220) | (0.0702) | (0.1126) | (0.2257) |
| Observations | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 24,055 | 11,911 | 29,401 |
| Weak identification test (Cragg- Donald Wald F statistic) | 638.1 | 638.1 | 638.1 | 638.1 | 638.1 | 546.8 | 243.1 | 638.1 |
| Weak identification test (Kleibergen- Paap rk Wald F statistic) | 134.9 | 134.9 | 134.9 | 134.9 | 134.9 | 119.8 | 66.0 | 134.9 |

Table 3: 2SLS regression of employment outcomes

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1.

| | U | 1 5 | | U | 1 1 | | | |
|-------------------------------------|------------|-------------|--------------------------|---------------|---------------|--------------|------------|------------|
| | | | | Depender | nt variables | | | |
| | Currently | Having wage | Self- | Self- | Having | Log of | Log of | Log of |
| | working | job (yes=1, | employed | employed | social | number of | monthly | monthly |
| Explanatory variables | (yes=1, | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (all |
| | no=0) | | work (yes=1, $n_{2}=0$) | (yes=1, na=0) | (yes = 1, no) | hours in the | workers) | workers) |
| | (1) | (2) | no=0) | no=0) | = 0) | last month | (7) | (0) |
| T C '4 C ' | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Log of per capita transfer income | 0.0024 | 0.0221*** | -0.0201*** | 0.0004 | -0.0018 | -0.0034 | -0.0168 | 0.1707** |
| | (0.0056) | (0.0081) | (0.0056) | (0.0075) | (0.0074) | (0.0123) | (0.0258) | (0.0694) |
| (Oldest's month age-960)*I{Oldest's | 0.0002 | -0.0001 | 0.0005 | -0.0001 | 0.0003 | -0.0004 | -0.0008 | -0.0013 |
| month age>=960} | (0.0003) | (0.0004) | (0.0003) | (0.0004) | (0.0004) | (0.0006) | (0.0015) | (0.0035) |
| (Oldest's month age-960) | -0.0001 | -0.0010*** | 0.0005*** | 0.0005** | -0.0003 | 0.0004 | 0.0007 | -0.0083*** |
| | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0003) | (0.0004) | (0.0009) | (0.0021) |
| Age | 0.0831*** | 0.0605*** | 0.0128*** | 0.0097*** | 0.0340*** | 0.0570*** | 0.0361*** | 0.5156*** |
| | (0.0013) | (0.0014) | (0.0010) | (0.0012) | (0.0010) | (0.0033) | (0.0056) | (0.0116) |
| Age squared | -0.0010*** | -0.0008*** | -0.0001*** | -0.0000*** | -0.0005*** | -0.0008*** | -0.0005*** | -0.0070*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0002) |
| Male (male=1, female=0) | 0.0521*** | 0.1424*** | -0.0217*** | -0.0687*** | -0.0187*** | 0.0961*** | 0.1733*** | 1.2431*** |
| | (0.0041) | (0.0060) | (0.0035) | (0.0046) | (0.0048) | (0.0070) | (0.0159) | (0.0513) |
| Ethnic minorities (yes=1, no=0) | 0.1413*** | -0.0944*** | -0.0821*** | 0.3178*** | -0.0999*** | -0.0131 | -0.6170*** | -0.9701*** |
| | (0.0053) | (0.0081) | (0.0043) | (0.0092) | (0.0054) | (0.0130) | (0.0510) | (0.0634) |
| Urban areas (urban=1, rural=0) | -0.0724*** | 0.1025*** | 0.0420*** | -0.2169*** | 0.1437*** | 0.1390*** | 0.3890*** | 1.0267*** |
| | (0.0071) | (0.0092) | (0.0079) | (0.0074) | (0.0093) | (0.0125) | (0.0267) | (0.0785) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7636*** | -0.7154*** | -0.1346*** | 0.0863*** | -0.4135*** | 4.0494*** | 7.3863*** | -6.2547*** |
| | (0.0244) | (0.0265) | (0.0179) | (0.0227) | (0.0217) | (0.0698) | (0.1107) | (0.2243) |
| Observations | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 24,055 | 11,911 | 29,401 |

Table 4. 2SLS regression of employment outcomes on log of the per capita transfer income

Robust standard errors in parentheses. Standard errors are clustered at age month of the eldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1. Source: Estimation from VHLSSs from 2012 to 2020.

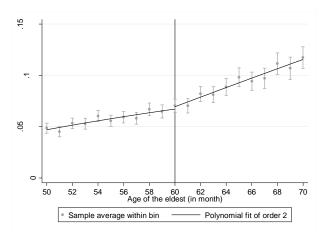
| | | The sample of p | eople aged 65-90 |) | | The sample of pe | ople aged 75-90 |) |
|--------------------------------------|------------|-----------------|------------------|------------------------|------------|------------------|-----------------|------------------------|
| | Currently | Self- | Self- | Log of | Currently | Self- | Self- | Log of |
| | working | employed | employed | number of | working | employed | employed | number of |
| Englander mensenisklag | (yes=1, | non-farm | farm work | working | (yes=1, | non-farm | farm work | working |
| Explanatory variables | no=0) | work (yes=1, | (yes=1, 0) | hours in the | no=0) | work (yes=1, | (yes=1, | hours in the |
| | | no=0) | no=0) | last month (working | | no=0) | no=0) | last month (working |
| | | | | people) | | | | people) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | -0.0882* | 0.0016 | -0.0822* | -0.0110 | -0.1084** | -0.0142 | -0.0862* | 0.0513 |
| | (0.0477) | (0.0210) | (0.0453) | (0.3078) | (0.0500) | (0.0188) | (0.0449) | (0.3155) |
| (Oldest's month age-960) * | 0.0009** | 0.0004*** | 0.0004 | 0.0019 | 0.0010*** | 0.0004*** | 0.0003 | 0.0010 |
| I{Oldest's month age>=960} | (0.0004) | (0.0002) | (0.0003) | (0.0019) | (0.0004) | (0.0001) | (0.0003) | (0.0022) |
| (Oldest's month age-960) | -0.0025*** | -0.0006*** | -0.0017*** | -0.0036*** | -0.0022*** | -0.0005*** | -0.0015*** | -0.0032*** |
| | (0.0002) | (0.0001) | (0.0002) | (0.0011) | (0.0003) | (0.0001) | (0.0002) | (0.0011) |
| Male (male=1, female=0) | 0.0318*** | -0.0091*** | 0.0304*** | -0.0166 | 0.0654*** | -0.0027 | 0.0575*** | 0.0506* |
| | (0.0054) | (0.0020) | (0.0050) | (0.0220) | (0.0065) | (0.0024) | (0.0060) | (0.0280) |
| Ethnic minorities (yes=1, no=0) | -0.0478*** | -0.0291*** | -0.0201 | 0.0699 | -0.0498*** | -0.0259*** | -0.0233* | 0.0496 |
| | (0.0148) | (0.0039) | (0.0146) | (0.0610) | (0.0139) | (0.0037) | (0.0138) | (0.0632) |
| Urban areas (urban=1, rural=0) | -0.1263*** | 0.0118** | -0.1441*** | 0.1746*** | -0.1272*** | 0.0026 | -0.1339*** | 0.1969*** |
| | (0.0094) | (0.0054) | (0.0092) | (0.0412) | (0.0091) | (0.0051) | (0.0085) | (0.0460) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | 0.1197*** | 0.0146* | 0.1061*** | 4.3534*** | 0.1063*** | 0.0230** | 0.0856*** | 4.2925*** |
| | (0.0140) | (0.0088) | (0.0131) | (0.1020) | (0.0151) | (0.0093) | (0.0132) | (0.1179) |
| Observations | 23,243 | 23,243 | 23,243 | 6,245 | 20,155 | 20,155 | 20,155 | 4,848 |

Table 5. 2SLS regression of employment outcomes of people aged from 65 and 75

Note: The sample includes individuals aged 65-90 living in households with the oldest member aged 75-84. The instrumental variable for 'Receiving the transfer' is I{Oldest's month age>=960} (dummy variable indicating the oldest aged from 80 years or 960 months). Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

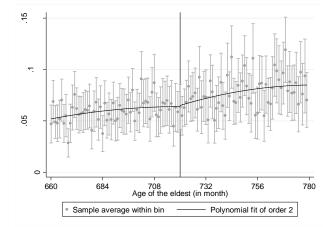
*** p<0.01, ** p<0.05, * p<0.1.

Figure A.1: The cash transfer receipt and age of the oldest household member with the threshold at 60 years



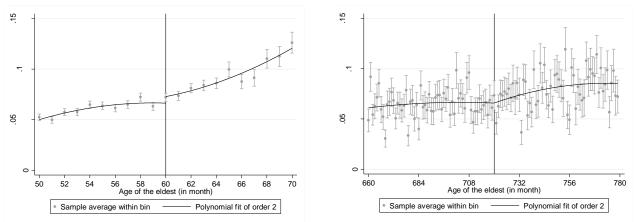
Panel A. Probability of households receiving the transfer and age in years of the oldest member

Panel A. Probability of individuals living in transferreceiving households and age in years of the oldest member



Panel B. Probability of households receiving the transfer and age in months of the oldest member

Panel B. Probability of individuals living in transferreceiving households and age in months of the oldest member



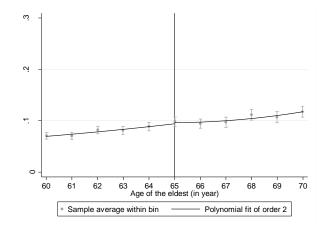
Note: Panels A and B graph the regression discontinuity plot of the probability of households receiving the cash transfers across age (in year and in month) of the oldest household member. The graph presents binned sample means at each age (in year and in month) and the 95% confidence interval. Panels A and B use household-level data and limit to the sample of households with the oldest household member aged 70 to 90 years.

Panels C and D show the regression discontinuity plot of the probability of individuals living in a household with the transfer across age (in year and in month) of the oldest household member. In these panels, the sample is limited to individuals (aged 15-64) living in a household with the oldest household member aged 70 to 90 years.

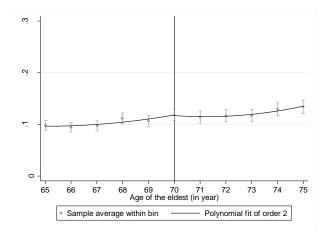
Figure A.2: The cash transfer receipt and age of the oldest household member at different

age thresholds

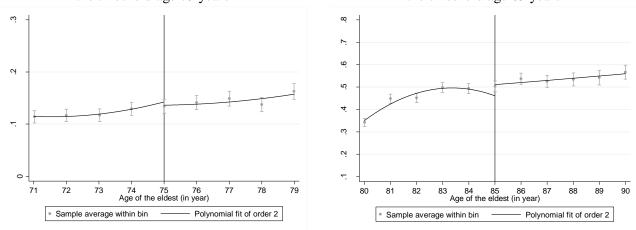
Panel A. Probability of households receiving the transfer and age in years of the oldest member around the threshold age 65 years



Panel C. Probability of households receiving the transfer and age in years of the oldest member around the threshold age 75 years Panel B. Probability of households receiving the transfer and age in years of the oldest member around the threshold age 70 years



Panel D. Probability of households receiving the transfer and age in years of the oldest member around the threshold age 85 years



Note: This figure graphs the regression discontinuity plot of the probability of households receiving the cash transfers across age (in year and in month) of the oldest household member. The graph presents binned sample means at each age (in year and in month) and the 95% confidence interval. This figure uses household-level data and limit to the sample of households with the oldest household member aged 60 to 90 years. Source: Estimation from VHLSSs 2012 to 2020.

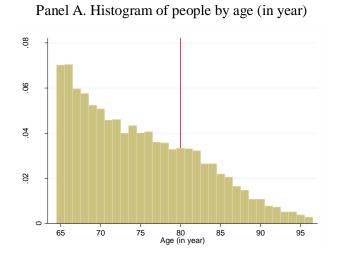
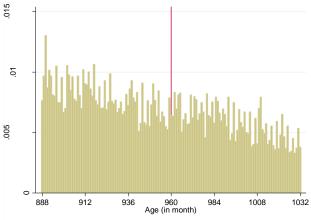
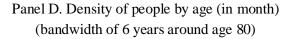
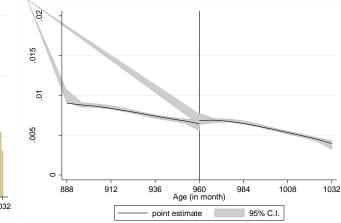


Figure A.3. The density of people by age in years and age in months

Panel C. Histogram of people by age (in month) (bandwidth of 6 years around age 80)







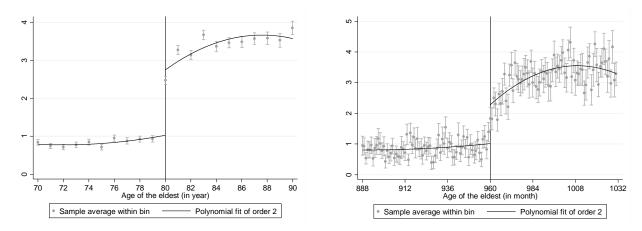
Note: This graph shows density of people around the threshold of aged 80 (or 960 months). Source: Estimation from VHLSSs 2012 to 2020.

Panel A. Density of people by age (in year)

Figure A.4. Log of per capita cash transfers and age of the oldest household member

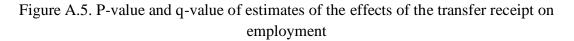
Panel A. Age in years of the oldest member Panel B.

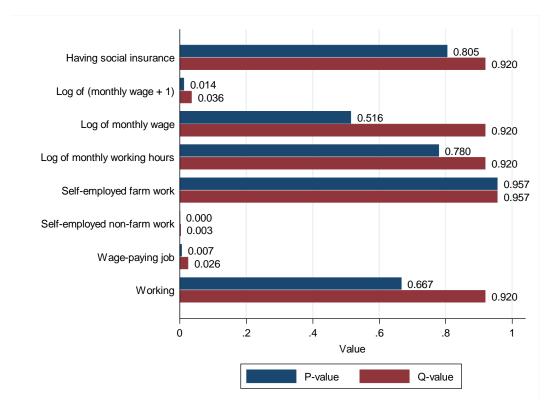
Panel B. Age in months of the oldest member



Note: This graph shows the regression discontinuity plot of log of per capita cash transfers across age (in year and in month) of the oldest household member. The graph presents binned sample means at each age month and the 95% confidence interval.

The sample is limited to individuals (aged 15-64) living in a household with the oldest household member aged 70 to 90 years.





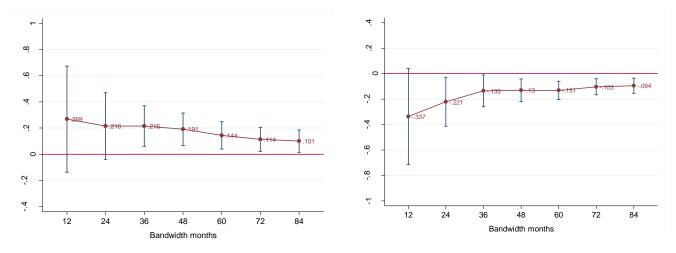
Note: The figure compares the p- and q-values of estimates of the transfer receipt on employment outcomes. The p-value is obtained from regressions reported in Table 3, while the q-value is computed using Simes' (1986) method.

Source: Authors' estimation from VHLSSs.

Figure A.6. Sensitivity of the estimated effect to bandwidth months

Panel A. Sensitivity of the estimated effect on the probability of having a wage job to bandwidth months

Panel B. Sensitivity of the estimated effect on the probability of having self-employed non-farm work to bandwidth months



Note: The figure graphs the point estimates and their 95% confidence intervals of the effect of the transfer receipt on the probability of having a wage job and the probability of non-farm work. The sample includes individuals aged 15-64 living in households with the oldest member in different age bandwidths. The estimates are from the 2SLS regression using the model specification similar to Table 3.

| Mariah laa | | All J | people aged 1 | 5-64 | | People age | d 15-64 living | g in household aged 75-84 | ds with at leas | st a member |
|-------------------------------------|--------------|--------------|---------------|--------------|--------------|--------------|----------------|------------------------------|-----------------|--------------|
| Variables | Year 2012 | Year 2014 | Year 2016 | Year 2018 | Year 2020 | Year 2012 | Year 2014 | Year 2016 | Year 2018 | Year 2020 |
| Proportion of working (%) | 81.2 | 82.0 | 82.1 | 81.9 | 81.9 | 81.2 | 81.9 | 81.4 | 80.1 | 80.5 |
| | (0.2) | (0.2) | (0.2) | (0.2) | (0.2) | (0.6) | (0.6) | (0.6) | (0.6) | (0.6) |
| Proportion of having a wage job | 37.8 | 38.5 | 40.4 | 42.6 | 47.4 | 38.9 | 39.8 | 42.1 | 43.4 | 47.4 |
| (%) | (0.3) | (0.3) | (0.3) | (0.3) | (0.3) | (0.8) | (0.9) | (0.9) | (1.1) | (0.9) |
| Proportion of having self-employed | 11.2 | 10.0 | 13.2 | 14.2 | 15.0 | 11.0 | 8.6 | 11.5 | 13.4 | 13.5 |
| non-farm work (%) | (0.2) | (0.2) | (0.2) | (0.2) | (0.3) | (0.6) | (0.5) | (0.6) | (0.6) | (0.7) |
| Proportion of having self-employed | 32.2 | 33.5 | 28.4 | 25.1 | 19.5 | 31.2 | 33.5 | 27.8 | 23.4 | 19.7 |
| farm work (%) | (0.5) | (0.5) | (0.5) | (0.5) | (0.4) | (1.0) | (0.9) | (1.0) | (1.4) | (0.9) |
| Proportion of having a job with | 13.7 | 14.0 | 16.1 | 17.8 | 22.4 | 14.6 | 15.6 | 18.0 | 20.0 | 22.2 |
| social insurance (%) | (0.3) | (0.3) | (0.3) | (0.3) | (0.4) | (0.7) | (0.7) | (0.8) | (0.8) | (0.9) |
| The number of monthly working | 183.4 | 186.1 | 185.7 | 187.4 | 193.0 | 185.1 | 185.4 | 189.5 | 189.8 | 189.7 |
| hours | (0.7) | (0.6) | (0.6) | (0.6) | (0.6) | (1.5) | (1.5) | (1.6) | (1.6) | (1.4) |
| Monthly wage (thousand VND, | 4208.3 | 4604.9 | 5212.0 | 5968.9 | 6634.1 | 4244.0 | 4798.4 | 5443.8 | 6304.3 | 6624.2 |
| computed for wage workers) | (49.4) | (48.0) | (52.8) | (54.5) | (65.7) | (101.7) | (124.8) | (113.9) | (163.7) | (154.4) |
| Monthly wage (thousand VND, | 1590.6 | 1772.4 | 2105.4 | 2540.1 | 3143.4 | 1651.7 | 1909.8 | 2290.7 | 2733.7 | 3141.3 |
| computed for all people) | (26.0) | (26.0) | (30.4) | (33.4) | (44.5) | (55.6) | (71.3) | (75.6) | (103.3) | (101.9) |
| Age | 36.5 | 37.5 | 38.2 | 39.0 | 38.8 | 35.8 | 36.3 | 36.9 | 37.6 | 38.0 |
| | (0.1) | (0.1) | (0.1) | (0.1) | (0.1) | (0.2) | (0.2) | (0.2) | (0.2) | (0.2) |
| Proportion of males (%) | 49.3 | 49.4 | 49.1 | 49.1 | 49.2 | 49.3 | 49.6 | 49.5 | 49.7 | 50.4 |
| | (0.1) | (0.1) | (0.1) | (0.1) | (0.1) | (0.5) | (0.5) | (0.5) | (0.5) | (0.6) |
| Proportion of ethnic minorities (%) | 14.1 | 14.9 | 15.6 | 15.6 | 13.3 | 14.7 | 16.3 | 16.3 | 15.8 | 14.4 |
| | (0.6) | (0.6) | (0.7) | (0.7) | (0.6) | (1.1) | (1.2) | (1.1) | (1.2) | (1.1) |
| Proportion of urban population (%) | 30.6 | 29.2 | 32.1 | 32.2 | 37.5 | 31.5 | 30.1 | 32.5 | 35.0 | 38.2 |
| | (1.0) | (0.9) | (1.0) | (1.0) | (1.1) | (1.6) | (1.6) | (1.7) | (1.7) | (1.7) |
| Number of schooling years | 8.6 | 8.8 | 8.9 | 9.1 | 9.4 | 8.9 | 9.0 | 9.1 | 9.3 | 9.6 |
| | (0.0) | (0.0) | (0.0) | (0.0) | (0.0) | (0.1) | (0.1) | (0.1) | (0.1) | (0.1) |
| Proportion of living in poor | 10.5 | 3.9 | 9.1 | 7.1 | 3.7 | 11.8 | 4.3 | 9.7 | 7.1 | 3.4 |
| households (%) | (0.3) | (0.2) | (0.4) | (0.3) | (0.2) | (0.8) | (0.5) | (0.8) | (0.7) | (0.4) |
| Number of observations | 124,419 | 119,803 | 117,850 | 113,689 | 113,274 | 6,265 | 6,164 | 6,018 | 5,743 | 5,211 |

Table A.1. Summary statistics of variables

Note: Monthly wages are adjusted to the 2020 January price using annual CPI. Poor households are those who are identified by local authorities as the poor. Standard errors of means in parentheses.

| | | Bandwidth of | of 60 months | | | Bandwidth of | of 36 months | |
|--------------------------------------|---------------------------|----------------------------|---------------------------|----------------------------|---------------------------|----------------------------|---------------------------|-------------------------------|
| Explanatory variables | Wage-paying job in public | Wage-paying job in private | Wage-paying job in public | Wage-paying job in private | Wage-paying job in public | Wage-paying job in private | Wage-paying job in public | Wage-paying job in private |
| | sector (yes = | sector (yes = |
| | 1, no = 0) | 1, no = 0) | 1, no = 0) | 1, no = 0) | 1, no = 0) | 1, no = 0) | 1, no = 0) | 1, no = 0) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | -0.0404 | 0.1842*** | | | -0.0643 | 0.2799*** | | |
| | (0.0306) | (0.0547) | | | (0.0458) | (0.0835) | | |
| Log of per capita transfer income | | | -0.0062 | 0.0283*** | | | -0.0101 | 0.0438*** |
| | | | (0.0047) | (0.0084) | | | (0.0071) | (0.0130) |
| (Oldest's month age-960)*I{Oldest's | 0.0005* | -0.0005 | 0.0005* | -0.0006 | -0.0004 | -0.0007 | -0.0003 | -0.0009 |
| month age>=960} | (0.0002) | (0.0004) | (0.0002) | (0.0004) | (0.0005) | (0.0008) | (0.0005) | (0.0008) |
| (Oldest's month age-960) | -0.0002 | -0.0008*** | -0.0002 | -0.0008*** | 0.0004 | -0.0016** | 0.0004 | -0.0016** |
| | (0.0002) | (0.0003) | (0.0002) | (0.0003) | (0.0004) | (0.0007) | (0.0004) | (0.0007) |
| Age | 0.0163*** | 0.0442*** | 0.0162*** | 0.0443*** | 0.0159*** | 0.0448*** | 0.0159*** | 0.0448*** |
| | (0.0008) | (0.0013) | (0.0008) | (0.0013) | (0.0010) | (0.0018) | (0.0010) | (0.0018) |
| Age squared | -0.0002*** | -0.0006*** | -0.0002*** | -0.0006*** | -0.0002*** | -0.0006*** | -0.0002*** | -0.0006*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) |
| Male (male=1, female=0) | 0.0072** | 0.1352*** | 0.0072** | 0.1352*** | 0.0097** | 0.1364*** | 0.0097** | 0.1364*** |
| | (0.0035) | (0.0061) | (0.0035) | (0.0061) | (0.0047) | (0.0081) | (0.0047) | (0.0081) |
| Ethnic minorities (yes=1, no=0) | -0.0201*** | -0.0775*** | -0.0210*** | -0.0734*** | -0.0235*** | -0.0736*** | -0.0249*** | -0.0677*** |
| | (0.0049) | (0.0085) | (0.0049) | (0.0085) | (0.0056) | (0.0112) | (0.0056) | (0.0112) |
| Urban areas (urban=1, rural=0) | 0.0808*** | 0.0212** | 0.0807*** | 0.0217** | 0.0749*** | 0.0319** | 0.0745*** | 0.0336*** |
| | (0.0067) | (0.0098) | (0.0068) | (0.0099) | (0.0081) | (0.0127) | (0.0081) | (0.0128) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.2125*** | -0.5043*** | -0.2128*** | -0.5026*** | -0.1908*** | -0.5263*** | -0.1916*** | -0.5228*** |
| | (0.0152) | (0.0249) | (0.0150) | (0.0246) | (0.0211) | (0.0346) | (0.0207) | (0.0339) |
| Observations | 29,401 | 29,401 | 29,401 | 29,401 | 17,867 | 17,867 | 17,867 | 17,867 |
| | | | | | | | 1 77 02 71 1 | |

Table A.2. 2SLS regression of wage job

Note: The sample includes individuals aged 15-64 living in households with the oldest member aged 75-84 and with the oldest member aged 77-82. The instrumental variable for 'Receiving the transfer' is I{Oldest's month age>=960} (dummy variable indicating the oldest aged from 80 years or 960 months). Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1.

| | | | | Depender | t variables | | | |
|-------------------------------------|----------------------|----------------------------|-----------------------|---------------------|-----------------------|----------------------------|---------------------|--------------------|
| | Currently working | Having wage job (yes=1, | Self- employed | Self- employed | Having social | Log of number of | Log of monthly | Log of monthly |
| Explanatory variables | (yes=1, 0) | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (all |
| | no=0) | | work (yes=1, no=0) | (yes=1, no=0) | (yes = 1, no) = 0) | hours in the last month | workers) | workers) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| | 0.0038 | 0.0349*** | -0.0317*** | 0.0006 | -0.0029 | -0.0056 | -0.0254 | 0.2696** |
| I{Oldest's month age>=960} | | | | | | | | |
| | (0.0089) 0.0002 | (0.0124) 0.0003 | (0.0082) 0.0001 | (0.0118) -0.0001 | (0.0117) 0.0002 | (0.0199) -0.0005 | (0.0389) -0.0011 | (0.1063) 0.0018 |
| (Oldest's month age-960)*I{Oldest's | | | | | | | | |
| month age>=960} | (0.0003) | (0.0003) | (0.0003) | (0.0003) | (0.0004) | (0.0005) | (0.0013) | (0.0030) |
| (Oldest's month age-960) | -0.0001 | -0.0010*** | 0.0005*** | 0.0005** | -0.0003 | 0.0004 | 0.0007 | -0.0079*** |
| | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0003) | (0.0003) | (0.0008) | (0.0020) |
| Age | 0.0831*** | 0.0604*** | 0.0129*** | 0.0097*** | 0.0340*** | 0.0571*** | 0.0369*** | 0.5145*** |
| | (0.0013) | (0.0013) | (0.0009) | (0.0012) | (0.0010) | (0.0033) | (0.0054) | (0.0115) |
| Age squared | -0.0010*** | -0.0008*** | -0.0001*** | -0.0000*** | -0.0005*** | -0.0008*** | -0.0005*** | -0.0070*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0002) |
| Male (male=1, female=0) | 0.0521*** | 0.1424*** | -0.0216*** | -0.0687*** | -0.0187*** | 0.0961*** | 0.1723*** | 1.2427*** |
| | (0.0041) | (0.0059) | (0.0034) | (0.0046) | (0.0048) | (0.0070) | (0.0161) | (0.0504) |
| Ethnic minorities (yes=1, no=0) | 0.1410*** | -0.0975*** | -0.0792*** | 0.3177*** | -0.0996*** | -0.0127 | -0.6134*** | -0.9945*** |
| | (0.0052) | (0.0085) | (0.0043) | (0.0093) | (0.0054) | (0.0130) | (0.0500) | (0.0655) |
| Urban areas (urban=1, rural=0) | -0.0740*** | 0.0880*** | 0.0552*** | -0.2171*** | 0.1449*** | 0.1411*** | 0.4011*** | 0.9146*** |
| | (0.0063) | (0.0082) | (0.0069) | (0.0071) | (0.0079) | (0.0096) | (0.0205) | (0.0692) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7622*** | -0.7020*** | -0.1467*** | 0.0866*** | -0.4146*** | 4.0453*** | 7.3625*** | -6.1515*** |
| | (0.0237) | (0.0246) | (0.0158) | (0.0215) | (0.0197) | (0.0663) | (0.0958) | (0.2094) |
| Observations | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 24,055 | 11,911 | 29,401 |
| R-squared | 0.311 | 0.115 | 0.051 | 0.187 | 0.096 | 0.065 | 0.139 | 0.126 |

Table A.3. Reduced-form regression of employment outcomes

Note: The sample includes individuals aged 15-64 living in households with the Oldest member aged 75-84. Robust standard errors in parentheses. Standard errors are clustered at age month of the Oldest and household levels. *** p<0.01, ** p<0.05, * p<0.1.

| | | | Depender | t variables | | |
|--------------------------------------|-------------------------------|------------|--------------------------------------------|-----------------------------------------|---------------------------------|--------------------|
| Explanatory variables | Male (male=1, female=0) | Age | Ethnic minorities (yes=1, Kinh=0) | Urban dummy (urban=1, rural=0) | Number of schooling years | Education level |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| Receiving the transfer (yes=1, no=0) | -0.0064 | -1.0582 | -0.0035 | -0.1316 | -0.5060 | -0.2248 |
| | (0.0257) | (1.6707) | (0.0587) | (0.0834) | (0.4645) | (0.1548) |
| (Oldest's month age-960)*I{Oldest's | -0.0001 | 0.0097 | -0.0005 | 0.0004 | 0.0047 | 0.0018* |
| month age>=960} | (0.0002) | (0.0120) | (0.0004) | (0.0006) | (0.0032) | (0.0010) |
| (Oldest's month age-960) | 0.0000 | 0.0207*** | 0.0001 | -0.0002 | -0.0038* | -0.0011 |
| | (0.0001) | (0.0072) | (0.0003) | (0.0004) | (0.0022) | (0.0008) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | 0.4409*** | 55.3824*** | 0.1377*** | 0.3620*** | 6.9997*** | 2.5623*** |
| | (0.0085) | (0.4778) | (0.0175) | (0.0247) | (0.1369) | (0.0490) |
| Observations | 52,644 | 52,644 | 52,644 | 52,644 | 52,644 | 52,644 |

Table A.4. 2SLS regression of employment outcomes on exogenous variables

Note: The sample includes individuals aged 15-84 living in households with the oldest member aged 75-84. The instrumental variable for 'Receiving the transfer' is I{Oldest's month age>=960} (dummy variable indicating the oldest aged from 80 years or 960 months). Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1.

| | | | Dependen | t variables | | |
|---------------------------------------|-------------------------------|------------|--------------------------------------------|-----------------------------------------|---------------------------------|--------------------|
| Explanatory variables | Male (male=1, female=0) | Age | Ethnic minorities (yes=1, Kinh=0) | Urban dummy (urban=1, rural=0) | Number of schooling years | Education level |
| | (1) | (2) | (3) | (4) | (5) | (6) |
| I{Oldest's month age>=960} | -0.0015 | -0.2492 | -0.0008 | -0.0310 | -0.1191 | -0.0529 |
| | (0.0061) | (0.3912) | (0.0138) | (0.0191) | (0.1117) | (0.0374) |
| (Oldest's month age-960) * I{Oldest's | -0.0001 | 0.0073 | -0.0005 | 0.0001 | 0.0036 | 0.0013 |
| month age>=960} | (0.0002) | (0.0106) | (0.0004) | (0.0006) | (0.0030) | (0.0010) |
| (Oldest's month age-960) | 0.0000 | 0.0200*** | 0.0001 | -0.0003 | -0.0042** | -0.0012 |
| | (0.0001) | (0.0065) | (0.0003) | (0.0003) | (0.0021) | (0.0008) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | 0.4399*** | 55.2155*** | 0.1371*** | 0.3412*** | 6.9199*** | 2.5268*** |
| | (0.0049) | (0.2625) | (0.0097) | (0.0132) | (0.0763) | (0.0280) |
| Observations | 52,644 | 52,644 | 52,644 | 52,644 | 52,644 | 52,644 |
| R-squared | 0.000 | 0.001 | 0.000 | 0.002 | 0.001 | 0.001 |

Table A.5. Reduced-form regression of employment outcomes on exogenous variables

Note: The sample includes individuals aged 15-84 living in households with the oldest member aged 75-84.

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1. Source: Estimation from VHLSSs from 2012 to 2020.

| | | | | Baseline depe | ndent variables | | | |
|-----------------------------------------------------------------------------|------------------------------------------|-------------------------------------|--------------------------------------------------------|----------------------------------------------------|----------------------------------------------------|--------------------------------------------------------------|---------------------------------------------|--------------------------------------------|
| Explanatory variables | Currently working (yes=1, no=0) | Having wage job (yes=1, no=0) | Self- employed non-farm work (yes=1, no=0) | Self- employed farm work (yes=1, no=0) | Having social insurance (yes = 1, no = 0) | Log of number of working hours in the last month | Log of monthly wage (wage workers) | Log of monthly wage (all workers) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| I{Oldest's endline month age>=960} | 0.0021 (0.0254) | -0.0124 (0.0413) | -0.0333 (0.0345) | 0.0477 (0.0338) | -0.0074 (0.0396) | -0.0192 (0.0436) | -0.0610 (0.0960) | -0.1084 (0.3312) |
| (Oldest's endline month age- 060)*U(Oldest's andline month age- | 0.0028** | -0.0002 | -0.0017 | 0.0047*** | 0.0010 | -0.0047** | -0.0010 | 0.0001 |
| 960)*I{Oldest's endline month age>=960} (Oldest's endline month age-960) | (0.0013) -0.0017** | (0.0019) -0.0007 | (0.0015) 0.0015 | (0.0017) -0.0025** | (0.0018) -0.0012 | (0.0022) 0.0012 | (0.0044) 0.0014 | (0.0155) -0.0095 |
| Age | (0.0007) 0.0844*** (0.0040) | (0.0016) 0.0449*** (0.0032) | (0.0010) 0.0170*** (0.0022) | (0.0012) 0.0225*** (0.0036) | (0.0014) 0.0280*** (0.0025) | (0.0016) 0.0813*** (0.0100) | (0.0023) 0.0251 (0.0158) | (0.0127) 0.4363*** (0.0255) |
| Age squared | -0.0010*** (0.0001) | -0.0006*** (0.0000) | -0.0002*** (0.0000) | -0.0002*** (0.0001) | -0.0004*** (0.0000) | -0.0011*** (0.0001) | -0.0004* (0.0002) | -0.0058*** (0.0004) |
| Male (male=1, female=0) | 0.0529*** (0.0117) | 0.1208*** (0.0119) | -0.0161** (0.0080) | -0.0519*** (0.0107) | 0.0152 | 0.0855*** (0.0185) | 0.2337*** (0.0320) | 1.2850*** (0.0940) |
| Ethnic minorities (yes=1, no=0) | 0.1348*** (0.0116) | -0.2008*** (0.0187) | -0.0914*** (0.0097) | 0.4270*** (0.0249) | -0.0779*** (0.0120) | 0.0352 | -0.4247*** (0.0690) | -1.0527*** (0.1457) |
| Urban areas (urban=1, rural=0) | -0.0808*** (0.0155) | (0.0187) 0.1649*** (0.0227) | (0.0097) 0.0548*** (0.0200) | -0.3004*** (0.0188) | (0.0120) 0.1761*** (0.0202) | (0.0370) 0.1110*** (0.0297) | (0.0090) 0.3957*** (0.0420) | (0.1437) 1.0857*** (0.1878) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.8106*** (0.0755) | -0.5035*** (0.0593) | -0.1745*** (0.0368) | -0.1326** (0.0602) | -0.3773*** (0.0502) | 3.6050*** (0.1969) | 7.2452*** (0.2824) | -5.1597*** (0.4836) |
| Observations R-squared | 3,978 0.325 | 3,978 0.126 | 3,978 0.052 | 3,978 0.255 | 3,978 0.099 | 3,321 0.083 | 1,521 0.204 | 3,978 0.125 |

Table A.6. Reduced-form regression of baseline employment outcomes (bandwidth of 24 months)

Note: The sample includes individuals aged 15-64 living in households with the oldest member aged 78-81 in the following survey after 2 years (so-called the endline survey). Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels. *** p<0.01, ** p<0.05, * p<0.1.

| | | | | Baseline depe | ndent variables | | | |
|-----------------------------------------|------------------------------------------|-------------------------------------|--------------------------------------------------------|----------------------------------------------------|-------------------------------------------------------|--------------------------------------------------------------|---------------------------------------------|--------------------------------------------|
| Explanatory variables | Currently working (yes=1, no=0) | Having wage job (yes=1, no=0) | Self- employed non-farm work (yes=1, no=0) | Self- employed farm work (yes=1, no=0) | Having social insurance (yes = 1, no = 0) | Log of number of working hours in the last month | Log of monthly wage (wage workers) | Log of monthly wage (all workers) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| I{Oldest's endline month age>=960} | -0.0111 | -0.0033 | -0.0284 | 0.0206 | -0.0070 | -0.0431 | -0.0064 | -0.1389 |
| | (0.0191) | (0.0290) | (0.0243) | (0.0273) | (0.0276) | (0.0360) | (0.0748) | (0.2354) |
| (Oldest's endline month age- | 0.0011** | -0.0004 | 0.0002 | 0.0013 | 0.0003 | -0.0017 | 0.0009 | -0.0027 |
| 960)*I{Oldest's endline month age>=960} | (0.0005) | (0.0009) | (0.0007) | (0.0008) | (0.0008) | (0.0012) | (0.0021) | (0.0070) |
| (Oldest's endline month age-960) | -0.0004 | -0.0007 | 0.0006 | -0.0003 | -0.0007 | 0.0007 | -0.0008 | -0.0061 |
| | (0.0003) | (0.0007) | (0.0005) | (0.0006) | (0.0006) | (0.0006) | (0.0012) | (0.0054) |
| Age | 0.0829*** | 0.0459*** | 0.0155*** | 0.0215*** | 0.0288*** | 0.0791*** | 0.0418*** | 0.4436*** |
| | (0.0029) | (0.0024) | (0.0016) | (0.0026) | (0.0019) | (0.0078) | (0.0115) | (0.0191) |
| Age squared | -0.0010*** | -0.0006*** | -0.0002*** | -0.0002*** | -0.0004*** | -0.0010*** | -0.0006*** | -0.0060*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0002) | (0.0003) |
| Male (male=1, female=0) | 0.0486*** | 0.1173*** | -0.0168*** | -0.0518*** | 0.0095 | 0.0827*** | 0.2130*** | 1.2825*** |
| | (0.0087) | (0.0098) | (0.0062) | (0.0093) | (0.0075) | (0.0142) | (0.0252) | (0.0786) |
| Ethnic minorities (yes=1, no=0) | 0.1234*** | -0.1790*** | -0.0702*** | 0.3727*** | -0.0737*** | 0.0407 | -0.3793*** | -1.0373*** |
| | (0.0095) | (0.0160) | (0.0086) | (0.0205) | (0.0092) | (0.0263) | (0.0533) | (0.1377) |
| Urban areas (urban=1, rural=0) | -0.0809*** | 0.1674*** | 0.0685*** | -0.3168*** | 0.1707*** | 0.1376*** | 0.4308*** | 1.1164*** |
| | (0.0114) | (0.0167) | (0.0137) | (0.0150) | (0.0147) | (0.0212) | (0.0349) | (0.1404) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7572*** | -0.5134*** | -0.1779*** | -0.0658 | -0.3829*** | 3.6402*** | 6.9084*** | -5.1777*** |
| | (0.0549) | (0.0484) | (0.0283) | (0.0509) | (0.0377) | (0.1547) | (0.2087) | (0.3819) |
| Observations | 6,670 | 6,670 | 6,670 | 6,670 | 6,670 | 5,572 | 2,558 | 6,670 |
| R-squared | 0.311 | 0.129 | 0.049 | 0.236 | 0.100 | 0.083 | 0.207 | 0.134 |

Table A.7. Reduced-form regression of baseline employment outcomes (bandwidth of 60 months)

Note: The sample includes individuals aged 15-64 living in households with the oldest member aged 75-84 in the following survey after 2 years (so-called the endline survey). Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels. *** p<0.01, ** p<0.05, * p<0.1.

| | e | 1 | 5 | e | 1 | | | | | | |
|-------------------------------------|------------------------------------------|-------------------------------------|--------------------------------------------------------|----------------------------------------------------|-------------------------------------------------------|--------------------------------------------------------------|---------------------------------------------|--------------------------------------------|--|--|--|
| | Dependent variables | | | | | | | | | | |
| Explanatory variables | Currently working (yes=1, no=0) | Having wage job (yes=1, no=0) | Self- employed non-farm work (yes=1, no=0) | Self- employed farm work (yes=1, no=0) | Having social insurance (yes = 1, no = 0) | Log of number of working hours in the last month | Log of monthly wage (wage workers) | Log of monthly wage (all workers) | | | |
| | (1) | (2) | (3) | (4) | (5) | last month (6) | (7) | (8) | | | |
| I{Oldest's month age>=960} | -0.0057 | -0.0061 | 0.0009 | 0.0039 | -0.0199 | -0.0007 | -0.0458 | -0.0575 | | | |
| (Ordest's month ages 900) | (0.0177) | (0.0220) | (0.0017) | (0.0035) | (0.0214) | (0.0359) | (0.0584) | (0.1664) | | | |
| (Oldest's month age-960)*I{Oldest's | 0.0012** | 0.0008 | 0.0000 | -0.0000 | 0.0001 | 0.0004 | 0.0010 | 0.0066 | | | |
| month age>=960} | (0.0012) | (0.0006) | (0.0000) | (0.0001) | (0.0006) | (0.0010) | (0.0016) | (0.0048) | | | |
| (Oldest's month age-960) | -0.0009** | -0.0008* | -0.0000 | -0.0000 | 0.0004 | -0.0008 | -0.0004 | -0.0063* | | | |
| (Ordest's month age 500) | (0.0004) | (0.0005) | (0.0000) | (0.0001) | (0.0005) | (0.0006) | (0.0011) | (0.0035) | | | |
| Age | 0.0336*** | 0.0417*** | 0.0009*** | 0.0016*** | 0.0236*** | 0.0923*** | 0.0459*** | 0.3333*** | | | |
| 50 | (0.0021) | (0.0022) | (0.0001) | (0.0002) | (0.0018) | (0.0054) | (0.0070) | (0.0172) | | | |
| Age squared | -0.0005*** | -0.0006*** | -0.0000*** | -0.0000*** | -0.0003*** | -0.0012*** | -0.0005*** | -0.0046*** | | | |
| -81 | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0001) | (0.0002) | | | |
| Male (male=1, female=0) | 0.1280*** | 0.1673*** | -0.0011* | -0.0044*** | 0.0015 | 0.0386*** | 0.1436*** | 1.3419*** | | | |
| | (0.0076) | (0.0084) | (0.0006) | (0.0007) | (0.0071) | (0.0128) | (0.0248) | (0.0648) | | | |
| Ethnic minorities (yes=1, no=0) | -0.0924*** | -0.0718*** | -0.0067*** | 0.0313*** | -0.0398*** | 0.0149 | -0.1823*** | -0.5850*** | | | |
| | (0.0123) | (0.0149) | (0.0009) | (0.0030) | (0.0088) | (0.0248) | (0.0415) | (0.1107) | | | |
| Urban areas (urban=1, rural=0) | 0.1848*** | 0.1396*** | 0.0052*** | -0.0248*** | 0.1632*** | 0.2707*** | 0.4521*** | 1.2598*** | | | |
| | (0.0140) | (0.0138) | (0.0010) | (0.0018) | (0.0130) | (0.0170) | (0.0328) | (0.1111) | | | |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | | | |
| Constant | -0.3848*** | -0.4524*** | -0.0184*** | -0.0279*** | -0.3042*** | 5.6170*** | 6.5679*** | -3.7770*** | | | |
| | (0.0367) | (0.0385) | (0.0024) | (0.0042) | (0.0310) | (0.1088) | (0.1265) | (0.2917) | | | |
| Observations | 28,401 | 28,401 | 28,401 | 28,401 | 6,090 | 21,754 | 8,769 | 28,401 | | | |
| R-squared | 0.179 | 0.101 | 0.117 | 0.375 | 0.088 | 0.507 | 0.194 | 0.112 | | | |

Table A.8. Reduced-form regression of employment outcomes using the sample of VHLSSs 2004 to 2010

Note: The sample includes individuals aged 15-64 living in households with the oldest member aged 75-84. Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels. *** p<0.01, ** p<0.05, * p<0.1.

| | | | | Depender | nt variables | | | |
|--------------------------------------|------------------------------------------|-------------------------------------|--------------------------------------------------------|----------------------------------------------------|-------------------------------------------------------|--------------------------------------------------------------|---------------------------------------------|--------------------------------------------|
| Explanatory variables | Currently working (yes=1, no=0) | Having wage job (yes=1, no=0) | Self- employed non-farm work (yes=1, no=0) | Self- employed farm work (yes=1, no=0) | Having social insurance (yes = 1, no = 0) | Log of number of working hours in the last month | Log of monthly wage (wage workers) | Log of monthly wage (all workers) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | 0.0122 | 0.1067** | -0.1356*** | 0.0411 | -0.0461 | -0.0705 | -0.2711 | 0.7721* |
| | (0.0478) | (0.0498) | (0.0349) | (0.0583) | (0.0522) | (0.0840) | (0.1857) | (0.4329) |
| (Oldest's month age-960)*I{Oldest's | 0.0001 | -0.0000 | 0.0005* | -0.0004 | 0.0004 | -0.0004 | 0.0001 | -0.0002 |
| month age>=960} | (0.0004) | (0.0004) | (0.0003) | (0.0005) | (0.0004) | (0.0006) | (0.0015) | (0.0033) |
| (Oldest's month age-960) | -0.0001 | -0.0013*** | 0.0005*** | 0.0006** | -0.0005* | 0.0001 | 0.0006 | -0.0107*** |
| | (0.0002) | (0.0002) | (0.0002) | (0.0003) | (0.0003) | (0.0004) | (0.0011) | (0.0022) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | 0.8075*** | 0.3620*** | 0.1346*** | 0.3109*** | 0.1493*** | 5.1269*** | 8.1873*** | 2.9634*** |
| | (0.0107) | (0.0122) | (0.0087) | (0.0141) | (0.0126) | (0.0218) | (0.0463) | (0.1031) |
| Observations | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 24,055 | 11,911 | 29,401 |

Table A.9. 2SLS regression of employment outcomes using the small model specification

Note: The sample includes individuals aged 15-64 living in households with the oldest member aged 75-84. The instrumental variable for 'Receiving the transfer' is I{Oldest's month age>=960} (dummy variable indicating the oldest aged from 80 years or 960 months). Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels. *** p < 0.01, ** p < 0.05, * p < 0.1. Source: Estimation from VHLSSs from 2012 to 2020.

| | U | 1 | 2 | U | U | 1 | | | | |
|---------------------------------------|---------------------|-------------|--------------|------------|---------------|--------------|------------|-----------|--|--|
| | Dependent variables | | | | | | | | | |
| | Currently | Having wage | Self- | Self- | Having | Log of | Log of | Log of | | |
| | working | job (yes=1, | employed | employed | social | number of | monthly | monthly | | |
| Explanatory variables | (yes=1, | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (al | | |
| | no=0) | | work (yes=1, | (yes=1, | (yes = 1, no) | hours in the | workers) | workers | | |
| | | | no=0) | no=0) | = 0) | last month | | | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | | |
| Receiving the transfer (yes=1, no=0) | 0.0133 | 0.1416*** | -0.1163*** | -0.0120 | 0.0113 | -0.0315 | 0.0317 | 1.1308** | | |
| | (0.0328) | (0.0477) | (0.0338) | (0.0396) | (0.0396) | (0.0755) | (0.1370) | (0.3940) | | |
| (Oldest's month age-960)*I{Oldest's | 0.0001 | -0.0004 | 0.0005* | -0.0000 | -0.0002 | -0.0002 | -0.0012 | -0.0035 | | |
| month age>=960} | (0.0003) | (0.0004) | (0.0003) | (0.0003) | (0.0003) | (0.0006) | (0.0012) | (0.0030) | | |
| (Oldest's month age-960) | -0.0000 | -0.0009*** | 0.0004** | 0.0004 *** | -0.0002 | 0.0002 | 0.0001 | -0.0068** | | |
| | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0003) | (0.0007) | (0.0019) | | |
| Age | 0.0798*** | 0.0510*** | 0.0164*** | 0.0124*** | 0.0229*** | 0.0586*** | 0.0324*** | 0.4299** | | |
| | (0.0013) | (0.0013) | (0.0010) | (0.0012) | (0.0009) | (0.0035) | (0.0056) | (0.0110) | | |
| Age squared | -0.0009*** | -0.0007*** | -0.0002*** | -0.0001*** | -0.0003*** | -0.0008*** | -0.0004*** | -0.0058* | | |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0001) | | |
| Male (male=1, female=0) | 0.0499*** | 0.1420*** | -0.0252*** | -0.0670*** | -0.0225*** | 0.0960*** | 0.2191*** | 1.2369** | | |
| | (0.0040) | (0.0059) | (0.0035) | (0.0044) | (0.0042) | (0.0069) | (0.0155) | (0.0507 | | |
| Ethnic minorities (yes=1, no=0) | 0.1410*** | 0.0292*** | -0.0660*** | 0.1779*** | -0.0176*** | -0.0406** | -0.3644*** | 0.1078 | | |
| | (0.0074) | (0.0112) | (0.0074) | (0.0122) | (0.0065) | (0.0173) | (0.0604) | (0.0843 | | |
| Urban areas (urban=1, rural=0) | -0.0705*** | 0.0159* | 0.0578*** | -0.1442*** | 0.0197*** | 0.0847*** | 0.1540*** | 0.1891** | | |
| | (0.0062) | (0.0085) | (0.0082) | (0.0077) | (0.0066) | (0.0115) | (0.0209) | (0.0711) | | |
| Household size | 0.0053*** | -0.0110*** | 0.0037* | 0.0126*** | 0.0037* | 0.0054 | 0.0040 | -0.0840** | | |
| | (0.0017) | (0.0031) | (0.0020) | (0.0027) | (0.0020) | (0.0036) | (0.0080) | (0.0251) | | |
| Proportion of children below 16 years | 0.0788*** | 0.1104*** | 0.0121 | -0.0438* | 0.1048*** | -0.0813*** | 0.0584 | 0.9419** | | |
| | (0.0160) | (0.0250) | (0.0196) | (0.0236) | (0.0161) | (0.0313) | (0.0731) | (0.2081) | | |
| Proportion of people from 65 years | -0.0439* | 0.0152 | -0.0239 | -0.0352 | 0.0548*** | -0.1936*** | -0.1458 | 0.0907 | | |
| | (0.0247) | (0.0310) | (0.0212) | (0.0287) | (0.0210) | (0.0428) | (0.0925) | (0.2519 | | |
| Poor households classified by local | 0.0369*** | 0.0499*** | -0.0193* | 0.0063 | -0.0178 | -0.0050 | -0.1032** | 0.3634* | | |
| authorities | (0.0127) | (0.0181) | (0.0106) | (0.0165) | (0.0109) | (0.0212) | (0.0486) | (0.1482 | | |
| Education level dummies | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | | |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | | |
| Province fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | | |
| Constant | -0.8934*** | -0.6118*** | -0.2459*** | -0.0356 | -0.4062*** | 4.0783*** | 7.3370*** | -5.3001** | | |
| | (0.0297) | (0.0411) | (0.0258) | (0.0327) | (0.0240) | (0.0780) | (0.1137) | (0.3387) | | |
| Observations | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 24,055 | 11,911 | 29,401 | | |

Table A.10. 2SLS regression of employment outcomes using the large model specification

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1.

| | | | | Depender | nt variables | | | |
|--------------------------------------|------------|-------------|--------------|------------|---------------|--------------|------------|-----------|
| | Currently | Having wage | Self- | Self- | Having | Log of | Log of | Log of |
| | working | job (yes=1, | employed | employed | social | number of | monthly | monthly |
| Explanatory variables | (yes=1, | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (all |
| | no=0) | | work (yes=1, | (yes=1, | (yes = 1, no) | hours in the | workers) | workers) |
| | | | no=0) | no=0) | = 0) | last month | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | 0.0486 | 0.1785*** | -0.1186*** | -0.0114 | 0.0267 | -0.0210 | 0.0753 | 1.4702*** |
| | (0.0452) | (0.0663) | (0.0442) | (0.0668) | (0.0548) | (0.1025) | (0.1576) | (0.5631) |
| (Oldest's month age-960)*I{Oldest's | 0.0001 | -0.0004 | 0.0005 | 0.0000 | -0.0001 | -0.0000 | -0.0001 | -0.0032 |
| month age>=960} | (0.0003) | (0.0004) | (0.0003) | (0.0005) | (0.0004) | (0.0007) | (0.0011) | (0.0038) |
| (Oldest's month age-960) | -0.0002 | -0.0011*** | 0.0004* | 0.0005* | -0.0004 | 0.0001 | -0.0002 | -0.0090** |
| | (0.0002) | (0.0003) | (0.0002) | (0.0003) | (0.0003) | (0.0005) | (0.0007) | (0.0023) |
| Age | 0.0818*** | 0.0570*** | 0.0128*** | 0.0120*** | 0.0309*** | 0.0582*** | 0.0389*** | 0.4860** |
| | (0.0015) | (0.0016) | (0.0011) | (0.0013) | (0.0011) | (0.0035) | (0.0053) | (0.0132) |
| Age squared | -0.0010*** | -0.0008*** | -0.0001*** | -0.0001*** | -0.0004*** | -0.0008*** | -0.0005*** | -0.0066** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0002) |
| Male (male=1, female=0) | 0.0524*** | 0.1481*** | -0.0216*** | -0.0742*** | -0.0151*** | 0.0983*** | 0.1801*** | 1.2933** |
| | (0.0047) | (0.0061) | (0.0036) | (0.0051) | (0.0052) | (0.0075) | (0.0146) | (0.0519) |
| Ethnic minorities (yes=1, no=0) | 0.1416*** | -0.1044*** | -0.0725*** | 0.3185*** | -0.0908*** | -0.0052 | -0.3751*** | -0.9704** |
| | (0.0060) | (0.0099) | (0.0047) | (0.0114) | (0.0057) | (0.0148) | (0.0305) | (0.0782) |
| Urban areas (urban=1, rural=0) | -0.0761*** | 0.0984*** | 0.0531*** | -0.2276*** | 0.1488*** | 0.1488*** | 0.3783*** | 0.9815** |
| | (0.0076) | (0.0100) | (0.0082) | (0.0091) | (0.0095) | (0.0134) | (0.0237) | (0.0846) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7440*** | -0.6592*** | -0.1389*** | 0.0540** | -0.3736*** | 4.0205*** | 7.2582*** | -5.8111** |
| | (0.0290) | (0.0297) | (0.0191) | (0.0257) | (0.0237) | (0.0736) | (0.1027) | (0.2487) |
| Observations | 24,190 | 24,190 | 24,190 | 24,190 | 24,190 | 19,794 | 9,540 | 24,190 |

Table A.11. 2SLS regression of employment outcomes using the sample without the 2020 VHLSS

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1.

| | U | 1 5 | | U | 1 | | | |
|--------------------------------------|------------|-------------|--------------|------------|---------------|--------------|------------|------------|
| | | | | Depender | t variables | | | |
| | Currently | Having wage | Self- | Self- | Having | Log of | Log of | Log of |
| | working | job (yes=1, | employed | employed | social | number of | monthly | monthly |
| Explanatory variables | (yes=1, | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (all |
| | no=0) | | work (yes=1, | (yes=1, | (yes = 1, no) | hours in the | workers) | workers) |
| | | | no=0) | no=0) | = 0) | last month | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | 0.0260 | 0.1120** | -0.1071*** | 0.0211 | -0.0062 | -0.0642 | -0.1312 | 0.8513* |
| | (0.0350) | (0.0559) | (0.0370) | (0.0509) | (0.0475) | (0.0775) | (0.1736) | (0.4831) |
| (Oldest's month age-960)*I{Oldest's | 0.0002 | 0.0001 | 0.0001 | -0.0000 | 0.0003 | -0.0002 | -0.0011 | 0.0005 |
| month age>=960} | (0.0003) | (0.0004) | (0.0003) | (0.0005) | (0.0004) | (0.0007) | (0.0017) | (0.0038) |
| (Oldest's month age-960) | -0.0001 | -0.0011*** | 0.0006*** | 0.0003 | -0.0004 | 0.0005 | 0.0010 | -0.0085*** |
| | (0.0002) | (0.0003) | (0.0002) | (0.0002) | (0.0003) | (0.0004) | (0.0010) | (0.0025) |
| Age | 0.0845*** | 0.0639*** | 0.0123*** | 0.0083*** | 0.0364*** | 0.0560*** | 0.0396*** | 0.5471*** |
| | (0.0014) | (0.0015) | (0.0011) | (0.0014) | (0.0011) | (0.0039) | (0.0066) | (0.0127) |
| Age squared | -0.0010*** | -0.0009*** | -0.0001*** | -0.0000 | -0.0005*** | -0.0007*** | -0.0006*** | -0.0074*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0002) |
| Male (male=1, female=0) | 0.0525*** | 0.1390*** | -0.0237*** | -0.0628*** | -0.0229*** | 0.0948*** | 0.1693*** | 1.2236*** |
| | (0.0047) | (0.0074) | (0.0039) | (0.0055) | (0.0053) | (0.0084) | (0.0173) | (0.0627) |
| Ethnic minorities (yes=1, no=0) | 0.1367*** | -0.0940*** | -0.0821*** | 0.3127*** | -0.1103*** | -0.0190 | -0.6617*** | -0.9948*** |
| - | (0.0054) | (0.0085) | (0.0044) | (0.0095) | (0.0063) | (0.0145) | (0.0597) | (0.0668) |
| Urban areas (urban=1, rural=0) | -0.0683*** | 0.0942*** | 0.0450*** | -0.2075*** | 0.1372*** | 0.1300*** | 0.3771*** | 0.9609*** |
| | (0.0072) | (0.0104) | (0.0084) | (0.0080) | (0.0101) | (0.0139) | (0.0293) | (0.0902) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.8210*** | -0.6985*** | -0.0977*** | -0.0248 | -0.3730*** | 4.1227*** | 7.5972*** | -6.0784*** |
| | (0.0302) | (0.0339) | (0.0243) | (0.0281) | (0.0283) | (0.0838) | (0.1367) | (0.2914) |
| Observations | 23,134 | 23,134 | 23,134 | 23,134 | 23,134 | 18,916 | 9,598 | 23,134 |

Table A.12. 2SLS regression of employment outcomes using the sample without the 2012 VHLSS

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1. Source: Estimation from VHLSSs from 2014 to 2020.

| | Dependent variables | | | | | | | | | |
|--------------------------------------|------------------------------------------|-------------------------------------|--------------------------------------------------------|----------------------------------------------------|--------------------------------------------------------|--------------------------------------------------------------|---------------------------------------------|--------------------------------------------|--|--|
| Explanatory variables | Currently working (yes=1, no=0) | Having wage job (yes=1, no=0) | Self- employed non-farm work (yes=1, no=0) | Self- employed farm work (yes=1, no=0) | Having social insurance (yes = 1, no) = 0) | Log of number of working hours in the last month | Log of monthly wage (wage workers) | Log of monthly wage (all workers) | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | | |
| Receiving the transfer (yes=1, no=0) | 0.0173 | 0.1436*** | -0.0979*** | -0.0284 | -0.0040 | -0.0106 | -0.0216 | 1.1413*** | | |
| | (0.0332) | (0.0447) | (0.0315) | (0.0476) | (0.0381) | (0.0715) | (0.1569) | (0.3742) | | |
| (Oldest's month age-960)*I{Oldest's | 0.0004 | 0.0001 | 0.0001 | 0.0001 | 0.0003 | -0.0005 | -0.0006 | 0.0010 | | |
| month age>=960} | (0.0003) | (0.0004) | (0.0003) | (0.0004) | (0.0003) | (0.0006) | (0.0014) | (0.0030) | | |
| (Oldest's month age-960) | -0.0002 | -0.0011*** | 0.0005*** | 0.0005* | -0.0004* | 0.0004 | 0.0003 | -0.0091*** | | |
| | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0003) | (0.0008) | (0.0019) | | |
| Age | 0.0808*** | 0.0582*** | 0.0128*** | 0.0098*** | 0.0314*** | 0.0594*** | 0.0402*** | 0.4933*** | | |
| | (0.0011) | (0.0012) | (0.0008) | (0.0011) | (0.0009) | (0.0030) | (0.0056) | (0.0096) | | |
| Age squared | -0.0010*** | -0.0008*** | -0.0001*** | -0.0000*** | -0.0004*** | -0.0008*** | -0.0006*** | -0.0067*** | | |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0001) | | |
| Male (male=1, female=0) | 0.0507*** | 0.1432*** | -0.0255*** | -0.0669*** | -0.0171*** | 0.0979*** | 0.1672*** | 1.2370*** | | |
| | (0.0036) | (0.0051) | (0.0029) | (0.0043) | (0.0040) | (0.0063) | (0.0147) | (0.0417) | | |
| Ethnic minorities (yes=1, no=0) | 0.1416*** | -0.1095*** | -0.0802*** | 0.3314*** | -0.0882*** | 0.0030 | -0.5444*** | -1.0591*** | | |
| | (0.0046) | (0.0076) | (0.0039) | (0.0086) | (0.0049) | (0.0118) | (0.0449) | (0.0593) | | |
| Urban areas (urban=1, rural=0) | -0.0658*** | 0.0957*** | 0.0536*** | -0.2150*** | 0.1344*** | 0.1323*** | 0.3651*** | 0.9490*** | | |
| | (0.0054) | (0.0076) | (0.0063) | (0.0070) | (0.0071) | (0.0102) | (0.0249) | (0.0634) | | |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | | |
| Constant | -0.7259*** | -0.6916*** | -0.1316*** | 0.0973*** | -0.3817*** | 4.0017*** | 7.2987*** | -5.9972*** | | |
| | (0.0219) | (0.0244) | (0.0159) | (0.0226) | (0.0179) | (0.0607) | (0.1166) | (0.2033) | | |
| Observations | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 24,055 | 11,911 | 29,401 | | |

Table A.13. 2SLS regression of employment outcomes using the sample without the sampling weight

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1. Source: Estimation from VHLSSs from 2012 to 2020.

| | | rrors are cluster onth of the old | | Standard erro | ors are clustered level | at the village | Standard errors are clustered at age in month of the oldest and village level | | | |
|---------------------------------|-------------------------------|--------------------------------------|--------------------------------|-------------------------------|-------------------------------|--------------------------------|----------------------------------------------------------------------------------|-------------------------------|--------------------------------|--|
| Explanatory variables | Having wage job (yes=1, | Self- employed non-farm | Log of monthly wage (all | Having wage job (yes=1, | Self- employed non-farm | Log of monthly wage (all | Having wage job (yes=1, | Self- employed non-farm | Log of monthly wage (all | |
| | no=0) | work (yes=1, no=0) | workers) | no=0) | work (yes=1, no=0) | workers) | no=0) | work (yes=1, no=0) | workers) | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | |
| Receiving the transfer (yes=1, | 0.1439*** | -0.1308*** | 1.1113** | 0.1439*** | -0.1308*** | 1.1113** | 0.1439** | -0.1308*** | 1.1113** | |
| no=0) | (0.0529) | (0.0368) | (0.4501) | (0.0524) | (0.0360) | (0.4467) | (0.0626) | (0.0426) | (0.5304) | |
| (Oldest's month age- | -0.0001 | 0.0004 | -0.0008 | -0.0001 | 0.0004 | -0.0008 | -0.0001 | 0.0004 | -0.0008 | |
| 960)*I{Oldest's month age>=960} | (0.0004) | (0.0003) | (0.0034) | (0.0004) | (0.0003) | (0.0034) | (0.0005) | (0.0003) | (0.0041) | |
| (Oldest's month age-960) | -0.0010*** | 0.0005*** | -0.0084*** | -0.0010*** | 0.0005*** | -0.0084*** | -0.0010*** | 0.0005** | -0.0084*** | |
| | (0.0002) | (0.0002) | (0.0021) | (0.0002) | (0.0002) | (0.0020) | (0.0003) | (0.0002) | (0.0025) | |
| Age | 0.0605*** | 0.0128*** | 0.5153*** | 0.0605*** | 0.0128*** | 0.5153*** | 0.0605*** | 0.0128*** | 0.5153*** | |
| | (0.0014) | (0.0010) | (0.0117) | (0.0014) | (0.0010) | (0.0120) | (0.0013) | (0.0009) | (0.0113) | |
| Age squared | -0.0008*** | -0.0001*** | -0.0070*** | -0.0008*** | -0.0001*** | -0.0070*** | -0.0008*** | -0.0001*** | -0.0070*** | |
| | (0.0000) | (0.0000) | (0.0002) | (0.0000) | (0.0000) | (0.0002) | (0.0000) | (0.0000) | (0.0002) | |
| Male (male=1, female=0) | 0.1424*** | -0.0216*** | 1.2425*** | 0.1424*** | -0.0216*** | 1.2425*** | 0.1424*** | -0.0216*** | 1.2425*** | |
| | (0.0060) | (0.0035) | (0.0512) | (0.0060) | (0.0035) | (0.0513) | (0.0056) | (0.0036) | (0.0471) | |
| Ethnic minorities (yes=1, no=0) | -0.0976*** | -0.0791*** | -0.9952*** | -0.0976*** | -0.0791*** | -0.9952*** | -0.0976*** | -0.0791*** | -0.9952*** | |
| Eurine minorities (yes=1, no=0) | (0.0082) | (0.0042) | (0.0639) | (0.0093) | (0.0044) | (0.0732) | (0.0104) | (0.0048) | (0.0845) | |
| Urban areas (urban=1, rural=0) | 0.1021*** | 0.0424*** | 1.0235*** | 0.1021*** | 0.0424*** | 1.0235*** | 0.1021*** | 0.0424*** | 1.0235*** | |
| Orban areas (urban=1, tura=0) | (0.0091) | (0.0078) | (0.0779) | (0.0096) | (0.0079) | (0.0822) | (0.0104) | (0.0076) | (0.0891) | |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | |
| Constant | -0.7167*** | -0.1334*** | -6.2649*** | -0.7167*** | -0.1334*** | -6.2649*** | -0.7167*** | -0.1334*** | -6.2649*** | |
| | (0.0267) | (0.0181) | (0.2257) | (0.0270) | (0.0186) | (0.2288) | (0.0274) | (0.0187) | (0.2306) | |
| Observations | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | 29,401 | |

Table A.14. 2SLS regression of employment outcomes using the sample with different error clustering

Robust standard errors in parentheses. *** p<0.01, ** p<0.05, * p<0.1. Source: Estimation from VHLSSs from 2012 to 2020.

| Estimators | Currently working (yes=1, no=0) | Having wage job (yes=1, no=0) | Self- employed non-farm work (yes=1, no=0) | Self- employed farm work (yes=1, no=0) | Having social insurance (yes = 1, no = 0) | Log of number of working hours in the last month | Log of monthly wage (wage workers) | Log of monthly wage (all workers) |
|----------------|------------------------------------------|-------------------------------------|--------------------------------------------------------|----------------------------------------------------|----------------------------------------------------|--------------------------------------------------------------|---------------------------------------------|--------------------------------------------|
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| | | | One common MSE | | | | | |
| Conventional | 0.0409 | 0.2308*** | -0.1814*** | -0.0349 | -0.0288 | 0.0292 | -0.1809 | 1.8069*** |
| | (0.0737) | (0.0749) | (0.0683) | (0.0687) | (0.0630) | (0.1400) | (0.2374) | (0.6237) |
| Bias-corrected | 0.0472 | 0.2606*** | -0.2011*** | -0.0447 | -0.0312 | 0.0403 | -0.2144 | 2.0373*** |
| | (0.0737) | (0.0749) | (0.0683) | (0.0687) | (0.0630) | (0.1400) | (0.2374) | (0.6237) |
| Robust | 0.0472 | 0.2606*** | -0.2011*** | -0.0447 | -0.0312 | 0.0403 | -0.2144 | 2.0373*** |
| | (0.0798) | (0.0814) | (0.0738) | (0.0784) | (0.0688) | (0.1512) | (0.2570) | (0.6742) |
| Bandwidth | 30.1 | 35.0 | 33.0 | 48.8 | 40.0 | 34.0 | 41.7 | 35.2 |
| Observations | 15183 | 17104 | 16231 | 23880 | 20018 | 13575 | 8247 | 17650 |
| | Panel B. C | One common MSE | -optimal bandwid | th selector for th | he sum of regressio | on estimates | | |
| Conventional | 0.0441 | 0.2165*** | -0.1721*** | -0.0320 | -0.0204 | 0.0337 | -0.1730 | 1.6924*** |
| | (0.0569) | (0.0622) | (0.0635) | (0.0644) | (0.0524) | (0.1342) | (0.2805) | (0.5326) |
| Bias-corrected | 0.0492 | 0.2559*** | -0.1991*** | -0.0437 | -0.0204 | 0.0490 | -0.2157 | 1.9984*** |
| | (0.0569) | (0.0622) | (0.0635) | (0.0644) | (0.0524) | (0.1342) | (0.2805) | (0.5326) |
| Robust | 0.0492 | 0.2559*** | -0.1991*** | -0.0437 | -0.0204 | 0.0490 | -0.2157 | 1.9984*** |
| | (0.0635) | (0.0699) | (0.0719) | (0.0768) | (0.0597) | (0.1478) | (0.3038) | (0.5987) |
| Bandwidth | 40.0 | 44.6 | 35.4 | 52.7 | 50.8 | 35.7 | 35.0 | 43.5 |
| Observations | 19478 | 22015 | 17650 | 25849 | 24966 | 14430 | 6871 | 21391 |

Table A.15. Non-parametric estimation of the impact of the cash transfer program on employment (using rdrobust 'command')

Note: We use the 'rdrobust' command Stata, which report three different procedures:

- Conventional RDD estimates with a conventional variance estimator.

- Bias-corrected RDD estimates with a conventional variance estimator.

- Bias-corrected RDD estimates with a robust variance estimator.

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels. *** p<0.01, ** p<0.05, * p<0.1.

| | | | | Dependen | t variables | | | |
|--------------------------------------|------------|-------------|--------------|------------|---------------|--------------|------------|------------|
| | Currently | Having wage | Self- | Self- | Having | Log of | Log of | Log of |
| | working | job (yes=1, | employed | employed | social | number of | monthly | monthly |
| Explanatory variables | (yes=1, | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (all |
| | no=0) | | work (yes=1, | (yes=1, | (yes = 1, no) | hours in the | workers) | workers) |
| | | | no=0) | no=0) | = 0) | last month | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | -0.0051 | 0.2228*** | -0.2125*** | -0.0079 | -0.0362 | 0.0716 | -0.3059 | 1.6938** |
| | (0.0592) | (0.0800) | (0.0676) | (0.0564) | (0.0645) | (0.1220) | (0.2030) | (0.6849) |
| (Oldest's month age-960)*I{Oldest's | -0.0006 | -0.0009 | -0.0000 | 0.0002 | -0.0002 | -0.0035** | 0.0003 | -0.0080 |
| month age>=960} | (0.0007) | (0.0009) | (0.0006) | (0.0006) | (0.0006) | (0.0014) | (0.0023) | (0.0073) |
| (Oldest's month age-960) | 0.0006 | -0.0014** | 0.0016*** | 0.0004 | 0.0001 | 0.0009 | 0.0019 | -0.0103** |
| | (0.0005) | (0.0006) | (0.0005) | (0.0003) | (0.0004) | (0.0008) | (0.0013) | (0.0051) |
| Age | 0.0833*** | 0.0604*** | 0.0122*** | 0.0097*** | 0.0334*** | 0.0591*** | 0.0293*** | 0.5129*** |
| | (0.0017) | (0.0019) | (0.0011) | (0.0013) | (0.0013) | (0.0045) | (0.0067) | (0.0158) |
| Age squared | -0.0010*** | -0.0008*** | -0.0001*** | -0.0000** | -0.0005*** | -0.0008*** | -0.0004*** | -0.0070*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0001) | (0.0002) |
| Male (male=1, female=0) | 0.0568*** | 0.1468*** | -0.0200*** | -0.0715*** | -0.0166*** | 0.0958*** | 0.1755*** | 1.2790*** |
| | (0.0062) | (0.0080) | (0.0051) | (0.0050) | (0.0060) | (0.0100) | (0.0201) | (0.0681) |
| Ethnic minorities (yes=1, no=0) | 0.1452*** | -0.0968*** | -0.0792*** | 0.3162*** | -0.1005*** | -0.0016 | -0.6414*** | -0.9913*** |
| | (0.0067) | (0.0106) | (0.0063) | (0.0103) | (0.0056) | (0.0194) | (0.0605) | (0.0831) |
| Urban areas (urban=1, rural=0) | -0.0754*** | 0.1072*** | 0.0275** | -0.2184*** | 0.1383*** | 0.1472*** | 0.3467*** | 1.0506*** |
| | (0.0102) | (0.0117) | (0.0111) | (0.0087) | (0.0111) | (0.0186) | (0.0293) | (0.1003) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7500*** | -0.7150*** | -0.1094*** | 0.0922*** | -0.3917*** | 4.0024*** | 7.5574*** | -6.2086*** |
| | (0.0334) | (0.0353) | (0.0249) | (0.0262) | (0.0275) | (0.0958) | (0.1335) | (0.3032) |
| Bandwidth (in months) | 30 | 35 | 33 | 49 | 40 | 34 | 42 | 35 |
| Observations | 15,183 | 17,650 | 16,619 | 24,468 | 20,018 | 13,973 | 8,422 | 17,650 |

Table A.16. 2SLS regression of employment outcomes using bandwidths, which are estimated from one common MSE-optimal bandwidth selector

Note: The sample includes individuals aged 15-64 living in households with the oldest member aged within the bandwidths around 960 months. The instrumental variable for 'Receiving the transfer' is I{Oldest's month age>=960} (dummy variable indicating the oldest aged from 80 years or 960 months). Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1.

| | | | | e | | | | |
|--------------------------------------|------------|-------------|--------------|------------|---------------|--------------|------------|-----------|
| | | | | Dependen | nt variables | | | |
| | Currently | Having wage | Self- | Self- | Having | Log of | Log of | Log of |
| | working | job (yes=1, | employed | employed | social | number of | monthly | monthly |
| Explanatory variables | (yes=1, | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (all |
| | no=0) | | work (yes=1, | (yes=1, | (yes = 1, no) | hours in the | workers) | workers) |
| | | | no=0) | no=0) | = 0) | last month | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | 0.0149 | 0.2100*** | -0.1545** | -0.0025 | -0.0338 | 0.0747 | -0.2774 | 1.6199*** |
| | (0.0495) | (0.0649) | (0.0645) | (0.0515) | (0.0539) | (0.1169) | (0.2377) | (0.5707) |
| (Eldest's month age-960)*I{Eldest's | 0.0004 | -0.0006 | 0.0002 | 0.0000 | 0.0000 | -0.0031** | -0.0000 | -0.0048 |
| month age>=960} | (0.0005) | (0.0006) | (0.0007) | (0.0005) | (0.0004) | (0.0012) | (0.0030) | (0.0056) |
| (Eldest's month age-960) | -0.0001 | -0.0014*** | 0.0008 | 0.0004* | -0.0000 | 0.0007 | 0.0017 | -0.0110** |
| | (0.0003) | (0.0003) | (0.0006) | (0.0002) | (0.0003) | (0.0007) | (0.0020) | (0.0030) |
| Age | 0.0831*** | 0.0605*** | 0.0127*** | 0.0097*** | 0.0331*** | 0.0595*** | 0.0265*** | 0.5115*** |
| | (0.0015) | (0.0017) | (0.0011) | (0.0013) | (0.0011) | (0.0046) | (0.0072) | (0.0143) |
| Age squared | -0.0010*** | -0.0008*** | -0.0001*** | -0.0000** | -0.0005*** | -0.0008*** | -0.0004*** | -0.0070** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0001) | (0.0002) |
| Male (male=1, female=0) | 0.0528*** | 0.1441*** | -0.0208*** | -0.0700*** | -0.0174*** | 0.0960*** | 0.1760*** | 1.2512*** |
| | (0.0052) | (0.0072) | (0.0048) | (0.0048) | (0.0053) | (0.0094) | (0.0220) | (0.0627) |
| Ethnic minorities (yes=1, no=0) | 0.1430*** | -0.0931*** | -0.0806*** | 0.3190*** | -0.0992*** | -0.0063 | -0.6340*** | -0.9730** |
| | (0.0064) | (0.0091) | (0.0056) | (0.0100) | (0.0054) | (0.0185) | (0.0657) | (0.0720) |
| Urban areas (urban=1, rural=0) | -0.0698*** | 0.1120*** | 0.0352*** | -0.2154*** | 0.1382*** | 0.1497*** | 0.3598*** | 1.0746*** |
| | (0.0089) | (0.0101) | (0.0109) | (0.0080) | (0.0100) | (0.0174) | (0.0336) | (0.0883) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | |
| Constant | -0.7654*** | -0.7292*** | -0.1323*** | 0.0871*** | -0.3916*** | 3.9906*** | 7.5908*** | -6.2600** |
| | (0.0287) | (0.0327) | (0.0256) | (0.0244) | (0.0240) | (0.0963) | (0.1548) | (0.2801) |
| Bandwidth (in months) | 40 | 45 | 35 | 53 | 51 | 36 | 35 | 43 |
| Observations | 20,018 | 22,461 | 17,650 | 26,343 | 25,394 | 14,791 | 7,103 | 22,015 |

Table A.17. 2SLS regression of employment outcomes using bandwidths, which are estimated from one common MSE-optimal bandwidth selector for the sum of regression estimates

Note: The sample includes individuals aged 15-64 living in households with the oldest member aged within the bandwidth 44 months around 960 months. The instrumental variable for 'Receiving the transfer' is I{Oldest's month age>=960} (dummy variable indicating the oldest aged from 80 years or 960 months). Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1.

| | U | 1 2 | | U | | | | |
|--------------------------------------|------------|-------------|--------------|------------|---------------|--------------|------------|------------|
| | | | | Dependen | t variables | | | |
| | Currently | Having wage | Self- | Self- | Having | Log of | Log of | Log of |
| | working | job (yes=1, | employed | employed | social | number of | monthly | monthly |
| Explanatory variables | (yes=1, | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (all |
| | no=0) | | work (yes=1, | (yes=1, | (yes = 1, no) | hours in the | workers) | workers) |
| | | | no=0) | no=0) | = 0) | last month | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | 0.0323 | 0.2156*** | -0.1346** | -0.0487 | -0.0363 | 0.0772 | -0.2838 | 1.6361** |
| | (0.0533) | (0.0786) | (0.0640) | (0.0759) | (0.0687) | (0.1164) | (0.2299) | (0.6727) |
| (Oldest's month age-960)*I{Oldest's | 0.0002 | -0.0010 | 0.0005 | 0.0008 | -0.0002 | -0.0032** | -0.0001 | -0.0088 |
| month age>=960} | (0.0006) | (0.0008) | (0.0007) | (0.0009) | (0.0007) | (0.0013) | (0.0029) | (0.0071) |
| (Oldest's month age-960) | -0.0002 | -0.0012** | 0.0004 | 0.0005 | 0.0001 | 0.0007 | 0.0018 | -0.0092* |
| | (0.0004) | (0.0006) | (0.0006) | (0.0005) | (0.0005) | (0.0007) | (0.0018) | (0.0048) |
| Age | 0.0837*** | 0.0606*** | 0.0128*** | 0.0102*** | 0.0341*** | 0.0589*** | 0.0275*** | 0.5151*** |
| | (0.0016) | (0.0018) | (0.0011) | (0.0016) | (0.0013) | (0.0046) | (0.0071) | (0.0158) |
| Age squared | -0.0010*** | -0.0008*** | -0.0001*** | -0.0001** | -0.0005*** | -0.0008*** | -0.0004*** | -0.0070*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0001) | (0.0002) |
| Male (male=1, female=0) | 0.0537*** | 0.1461*** | -0.0205*** | -0.0719*** | -0.0182*** | 0.0961*** | 0.1783*** | 1.2745*** |
| | (0.0057) | (0.0079) | (0.0048) | (0.0059) | (0.0066) | (0.0096) | (0.0218) | (0.0673) |
| Ethnic minorities (yes=1, no=0) | 0.1395*** | -0.0971*** | -0.0822*** | 0.3188*** | -0.1008*** | -0.0051 | -0.6354*** | -0.9938*** |
| | (0.0065) | (0.0104) | (0.0056) | (0.0125) | (0.0060) | (0.0187) | (0.0650) | (0.0818) |
| Urban areas (urban=1, rural=0) | -0.0733*** | 0.1068*** | 0.0374*** | -0.2176*** | 0.1423*** | 0.1487*** | 0.3590*** | 1.0479*** |
| | (0.0096) | (0.0114) | (0.0107) | (0.0113) | (0.0118) | (0.0176) | (0.0328) | (0.0981) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7709*** | -0.7171*** | -0.1387*** | 0.0849*** | -0.4031*** | 4.0043*** | 7.5741*** | -6.2280*** |
| | (0.0307) | (0.0346) | (0.0251) | (0.0315) | (0.0294) | (0.0963) | (0.1508) | (0.2970) |
| Observations | 17,867 | 17,867 | 17,867 | 17,867 | 17,867 | 14,618 | 7,196 | 17,867 |

Table A.18. 2SLS regression of employment outcomes using the bandwidth of 36 months (3 years)

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1. Source: Estimation from VHLSSs from 2012 to 2020.

| | U | 1 2 | | U | | | | |
|--------------------------------------|------------|-------------|--------------|------------|---------------|--------------|------------|------------|
| | | | | Depender | nt variables | | | |
| | Currently | Having wage | Self- | Self- | Having | Log of | Log of | Log of |
| | working | job (yes=1, | employed | employed | social | number of | monthly | monthly |
| Explanatory variables | (yes=1, | no=0) | non-farm | farm work | insurance | working | wage (wage | wage (all |
| | no=0) | | work (yes=1, | (yes=1, | (yes = 1, no) | hours in the | workers) | workers) |
| | | | no=0) | no=0) | = 0) | last month | | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | 0.0185 | 0.1008** | -0.0944*** | 0.0121 | -0.0057 | 0.0036 | -0.0709 | 0.7761** |
| | (0.0302) | (0.0437) | (0.0306) | (0.0408) | (0.0370) | (0.0655) | (0.1308) | (0.3690) |
| (Oldest's month age-960)*I{Oldest's | 0.0004** | 0.0004 | 0.0001 | -0.0000 | 0.0004 | -0.0004 | -0.0004 | 0.0034 |
| month age>=960} | (0.0002) | (0.0003) | (0.0002) | (0.0003) | (0.0002) | (0.0004) | (0.0008) | (0.0023) |
| (Oldest's month age-960) | -0.0002 | -0.0010*** | 0.0005*** | 0.0003* | -0.0004** | 0.0002 | 0.0003 | -0.0081*** |
| | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0002) | (0.0003) | (0.0007) | (0.0018) |
| Age | 0.0827*** | 0.0599*** | 0.0144*** | 0.0084*** | 0.0336*** | 0.0600*** | 0.0413*** | 0.5109*** |
| | (0.0011) | (0.0012) | (0.0008) | (0.0011) | (0.0009) | (0.0029) | (0.0049) | (0.0103) |
| Age squared | -0.0010*** | -0.0008*** | -0.0002*** | -0.0000 | -0.0005*** | -0.0008*** | -0.0006*** | -0.0069*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0001) | (0.0001) |
| Male (male=1, female=0) | 0.0502*** | 0.1425*** | -0.0212*** | -0.0711*** | -0.0185*** | 0.1017*** | 0.1743*** | 1.2431*** |
| | (0.0037) | (0.0051) | (0.0029) | (0.0040) | (0.0041) | (0.0062) | (0.0141) | (0.0433) |
| Ethnic minorities (yes=1, no=0) | 0.1350*** | -0.0992*** | -0.0762*** | 0.3103*** | -0.1021*** | -0.0037 | -0.6000*** | -1.0061*** |
| - | (0.0049) | (0.0077) | (0.0037) | (0.0084) | (0.0051) | (0.0113) | (0.0444) | (0.0606) |
| Urban areas (urban=1, rural=0) | -0.0752*** | 0.0951*** | 0.0488*** | -0.2191*** | 0.1388*** | 0.1493*** | 0.3807*** | 0.9616*** |
| | (0.0059) | (0.0082) | (0.0067) | (0.0064) | (0.0079) | (0.0106) | (0.0219) | (0.0698) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7628*** | -0.6972*** | -0.1661*** | 0.1004*** | -0.4063*** | 3.9921*** | 7.2946*** | -6.1181*** |
| | (0.0214) | (0.0248) | (0.0157) | (0.0217) | (0.0186) | (0.0620) | (0.0943) | (0.2075) |
| Observations | 37,097 | 37,097 | 37,097 | 37,097 | 37,097 | 30,362 | 14,982 | 37,097 |

Table A.19. 2SLS regression of employment outcomes using the bandwidth of 84 months (7 years)

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1. Source: Estimation from VHLSSs from 2012 to 2020.

| | | e | 1 0 | | e | 1 | | |
|--------------------------------------|------------------------------|--------------------------|------------------------------|--------------------------|------------------------------|-------------------------------|------------------------------|----------------------|
| | | | | Dependen | t variables | | | |
| | Excluding individuals within | | Excluding individuals within | | Excluding individuals within | | Excluding individuals within | |
| | a bandwidth | of 3 months | a bandwidth | of 6 months | a bandwidth | n of 9 months | a bandwidth | of 12 months |
| | around the | e threshold | around the | e threshold | | e threshold | | e threshold |
| Explanatory variables | Having wage | Self- | Having wage | Self- | Having wage | Self- | Having wage | Self- |
| | job (yes=1, | employed | job (yes=1, | employed | job (yes=1, | employed | job (yes=1, | employed |
| | no=0) | non-farm | no=0) | non-farm | no=0) | non-farm | no=0) | non-farm |
| | | work (yes=1, $n_{2}=0$) | | work (yes=1, $n_{2}=0$) | | work (yes=1, $n_{\alpha}=0$) | | work (yes= n_{0}) |
| | (1) | no=0) | (2) | no=0) | (5) | no=0) | (7) | no=0) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Receiving the transfer (yes=1, no=0) | 0.1115** | -0.1041*** | 0.1393*** | -0.0926*** | 0.1244** | -0.1123*** | 0.1112* | -0.0834** |
| | (0.0543) | (0.0358) | (0.0505) | (0.0325) | (0.0580) | (0.0335) | (0.0669) | (0.0394) |
| (Oldest's month age-960)*I{Oldest's | 0.0000 | 0.0005 | -0.0000 | 0.0002 | -0.0003 | 0.0005* | -0.0003 | 0.0003 |
| month age>=960} | (0.0004) | (0.0003) | (0.0004) | (0.0003) | (0.0004) | (0.0003) | (0.0005) | (0.0003) |
| (Oldest's month age-960) | -0.0010*** | 0.0004* | -0.0011*** | 0.0004** | -0.0009*** | 0.0004* | -0.0008** | 0.0004 |
| | (0.0003) | (0.0002) | (0.0003) | (0.0002) | (0.0003) | (0.0002) | (0.0004) | (0.0003) |
| Age | 0.0606*** | 0.0130*** | 0.0604*** | 0.0131*** | 0.0607*** | 0.0129*** | 0.0604*** | 0.0131*** |
| | (0.0014) | (0.0010) | (0.0014) | (0.0010) | (0.0015) | (0.0010) | (0.0015) | (0.0011) |
| Age squared | -0.0008*** | -0.0001*** | -0.0008*** | -0.0001*** | -0.0008*** | -0.0001*** | -0.0008*** | -0.0001*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) |
| Male (male=1, female=0) | 0.1400*** | -0.0214*** | 0.1388*** | -0.0215*** | 0.1376*** | -0.0215*** | 0.1387*** | -0.0227*** |
| | (0.0061) | (0.0037) | (0.0063) | (0.0038) | (0.0066) | (0.0040) | (0.0069) | (0.0040) |
| Ethnic minorities (yes=1, no=0) | -0.0966*** | -0.0790*** | -0.0984*** | -0.0814*** | -0.0979*** | -0.0813*** | -0.0946*** | -0.0815** |
| | (0.0088) | (0.0043) | (0.0087) | (0.0043) | (0.0092) | (0.0044) | (0.0096) | (0.0045) |
| Urban areas (urban=1, rural=0) | 0.1018*** | 0.0447*** | 0.1060*** | 0.0431*** | 0.1042*** | 0.0445*** | 0.1010*** | 0.0477*** |
| | (0.0096) | (0.0076) | (0.0098) | (0.0077) | (0.0102) | (0.0077) | (0.0108) | (0.0081) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7112*** | -0.1427*** | -0.7143*** | -0.1412*** | -0.7132*** | -0.1379*** | -0.7069*** | -0.1438** |
| | (0.0272) | (0.0179) | (0.0280) | (0.0186) | (0.0298) | (0.0197) | (0.0326) | (0.0216) |
| Observations | 27,611 | 27,611 | 26,115 | 26,115 | 24,593 | 24,593 | 23,041 | 23,041 |

Table A.20. 2SLS regression of employment outcomes using 'donut' samples

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1.

| | • | 1 0 | | 1 1 0 | | | | | |
|--------------------------------------|------------|-----------------|-----------------|--------------|----------------------------------|--------------|------------|--------------|--|
| | | Bandwidth of 36 | months (3 years |) | Bandwidth of 84 months (7 years) | | | | |
| | Currently | Self- | Self- | Log of | Currently | Self- | Self- | Log of | |
| | working | employed | employed | number of | working | employed | employed | number of | |
| Explanatory variables | (yes=1, | non-farm | farm work | working | (yes=1, | non-farm | farm work | working | |
| | no=0) | work (yes=1, | (yes=1, | hours in the | no=0) | work (yes=1, | (yes=1, | hours in the | |
| | | no=0) | no=0) | last month | | no=0) | no=0) | last month | |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | |
| Receiving the transfer (yes=1, no=0) | -0.1251* | -0.0177 | -0.1039 | 0.1981 | -0.1137*** | -0.0122 | -0.0952** | 0.0828 | |
| | (0.0691) | (0.0242) | (0.0642) | (0.3988) | (0.0425) | (0.0161) | (0.0389) | (0.2552) | |
| (Oldest's month age-960) * | 0.0008 | 0.0002 | 0.0002 | -0.0019 | 0.0011*** | 0.0004*** | 0.0004* | 0.0007 | |
| I{Oldest's month age>=960} | (0.0008) | (0.0002) | (0.0007) | (0.0048) | (0.0003) | (0.0001) | (0.0003) | (0.0014) | |
| (Oldest's month age-960) | -0.0020*** | -0.0004* | -0.0012** | -0.0032* | -0.0022*** | -0.0005*** | -0.0015*** | -0.0032*** | |
| | (0.0006) | (0.0002) | (0.0005) | (0.0018) | (0.0002) | (0.0001) | (0.0002) | (0.0011) | |
| Male (male=1, female=0) | 0.0583*** | -0.0039 | 0.0537*** | 0.0349 | 0.0612*** | -0.0027 | 0.0546*** | 0.0572** | |
| | (0.0088) | (0.0031) | (0.0081) | (0.0350) | (0.0060) | (0.0021) | (0.0055) | (0.0263) | |
| Ethnic minorities (yes=1, no=0) | -0.0607*** | -0.0251*** | -0.0377** | 0.0595 | -0.0427*** | -0.0239*** | -0.0172 | 0.0482 | |
| | (0.0185) | (0.0042) | (0.0183) | (0.0815) | (0.0127) | (0.0034) | (0.0126) | (0.0626) | |
| Urban areas (urban=1, rural=0) | -0.1247*** | 0.0010 | -0.1282*** | 0.2442*** | -0.1195*** | 0.0030 | -0.1262*** | 0.2031*** | |
| | (0.0130) | (0.0065) | (0.0114) | (0.0602) | (0.0083) | (0.0045) | (0.0078) | (0.0441) | |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes | |
| Constant | 0.1269*** | 0.0340*** | 0.0956*** | 4.3705*** | 0.1081*** | 0.0237*** | 0.0855*** | 4.3053*** | |
| | (0.0198) | (0.0107) | (0.0191) | (0.1563) | (0.0139) | (0.0085) | (0.0121) | (0.1161) | |
| Observations | 12,447 | 12,447 | 12,447 | 2,890 | 22,917 | 22,917 | 22,917 | 5,178 | |

Table A.21. 2SLS regression of employment outcomes of people aged from 75 with different bandwidths

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels.

*** p<0.01, ** p<0.05, * p<0.1. Source: Estimation from VHLSSs from 2012 to 2020.

| υ | | 1 2 | \mathcal{O} | | 1 1 | | | |
|---------------------------------|------------------------------------------------------|--------------|-----------------------------------------------------|--------------|------------------------------------------------------|--------------|--------------------------------------------------------|--------------|
| | Individuals living in households without children | | Individuals living in households with children aged | | Males living in households with children aged 0-5 | | Females living in households with children aged 0-5 | |
| | | 10-5 | | -5 | | | | |
| | Having wage | Self- | Having wage | Self- | Having wage | Self- | Having wage | Self- |
| Explanatory variables | job (yes=1, | employed | job (yes=1, | employed | job (yes=1, | employed | job (yes=1, | employed |
| | no=0) | non-farm | no=0) | non-farm | no=0) | non-farm | no=0) | non-farm |
| | | work (yes=1, | | work (yes=1, | | work (yes=1, | | work (yes=1, |
| | (1) | no=0) | (2) | <u>no=0)</u> | (5) | <u>no=0)</u> | (7) | no=0) |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| Proportion of household | 0.0276*** | -0.0667*** | 0.0859*** | -0.0852*** | -0.0123 | -0.0866*** | 0.1758*** | -0.0847*** |
| members aged from 65 | (0.0070) | (0.0051) | (0.0115) | (0.0082) | (0.0149) | (0.0101) | (0.0138) | (0.0100) |
| Age | 0.0656*** | 0.0182*** | 0.0571*** | 0.0174*** | 0.0625*** | 0.0159*** | 0.0528*** | 0.0188*** |
| | (0.0004) | (0.0003) | (0.0006) | (0.0004) | (0.0008) | (0.0005) | (0.0007) | (0.0005) |
| Age squared | -0.0009*** | -0.0002*** | -0.0008*** | -0.0002*** | -0.0009*** | -0.0002*** | -0.0008*** | -0.0002*** |
| | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) | (0.0000) |
| Male (male=1, female=0) | 0.1465*** | -0.0297*** | 0.1687*** | -0.0168*** | | | | |
| | (0.0016) | (0.0011) | (0.0019) | (0.0013) | | | | |
| Ethnic minorities (yes=1, no=0) | -0.0707*** | -0.0904*** | -0.1073*** | -0.1002*** | -0.0932*** | -0.0990*** | -0.1183*** | -0.1021*** |
| | (0.0030) | (0.0014) | (0.0033) | (0.0016) | (0.0042) | (0.0019) | (0.0037) | (0.0019) |
| Urban areas (urban=1, rural=0) | 0.0767*** | 0.0602*** | 0.0910*** | 0.0512*** | 0.0799*** | 0.0428*** | 0.1012*** | 0.0588*** |
| | (0.0023) | (0.0019) | (0.0030) | (0.0025) | (0.0038) | (0.0030) | (0.0036) | (0.0029) |
| Year fixed effects | Yes | Yes | Yes | Yes | Yes | Yes | Yes | Yes |
| Constant | -0.7569*** | -0.2429*** | -0.5313*** | -0.2034*** | -0.4680*** | -0.1865*** | -0.4506*** | -0.2321*** |
| | (0.0068) | (0.0047) | (0.0101) | (0.0065) | (0.0147) | (0.0083) | (0.0125) | (0.0082) |
| Observations | 332,334 | 332,334 | 228,324 | 228,324 | 109,817 | 109,817 | 118,507 | 118,507 |
| R-squared | 0.117 | 0.060 | 0.139 | 0.044 | 0.099 | 0.039 | 0.126 | 0.048 |

Table A.22. OLS regressions of employment individuals aged 15-60 on the proportion of older members in households

Robust standard errors in parentheses. Standard errors are clustered at the household level. *** p<0.01, ** p<0.05, * p<0.1.

| | Dependent variable is school enrolment of children (yes=1, no=0) | | | | | | | |
|--------------------------------------|------------------------------------------------------------------|-----------------------------|------------------------------|------------------------------|--|--|--|--|
| Explanatory variables | Children aged 1-5 years | Children aged 6-10 years | Children aged 11-14 years | Children aged 15-17 years | | | | |
| | (1) | (2) | (3) | (4) | | | | |
| | 0.1538 | 0.0476 | 0.0451 | 0.2461 | | | | |
| Receiving the transfer (yes=1, no=0) | (0.1266) | (0.0421) | (0.0801) | (0.2144) | | | | |
| (Oldest's month age-960)*I{Oldest's | -0.0015 | -0.0010*** | -0.0004 | 0.0010 | | | | |
| month age>=960} | (0.0012) | (0.0004) | (0.0005) | (0.0014) | | | | |
| (Oldest's month age-960) | 0.0002 | 0.0002 | 0.0002 | -0.0019** | | | | |
| | (0.0006) | (0.0002) | (0.0004) | (0.0009) | | | | |
| Age | 0.3171*** | 0.1272*** | 0.2315** | 0.3777 | | | | |
| | (0.0285) | (0.0346) | (0.0960) | (0.5525) | | | | |
| Age squared | -0.0186*** | -0.0073*** | -0.0099** | -0.0139 | | | | |
| | (0.0048) | (0.0021) | (0.0039) | (0.0173) | | | | |
| Male (male=1, female=0) | 0.0131 | 0.0017 | -0.0122 | -0.0528*** | | | | |
| | (0.0144) | (0.0066) | (0.0080) | (0.0152) | | | | |
| | -0.0789*** | -0.0129 | -0.1197*** | -0.2454*** | | | | |
| Ethnic minorities (yes=1, no=0) | (0.0260) | (0.0185) | (0.0248) | (0.0360) | | | | |
| | 0.0821*** | 0.0010 | -0.0023 | 0.1079*** | | | | |
| Urban areas (urban=1, rural=0) | (0.0223) | (0.0085) | (0.0128) | (0.0319) | | | | |
| Year fixed effects | Yes | Yes | Yes | Yes | | | | |
| Constant | -0.3682*** | 0.4393*** | -0.3736 | -1.6930 | | | | |
| | (0.0519) | (0.1419) | (0.5924) | (4.4032) | | | | |
| Observations | 3,402 | 3,890 | 3,462 | 2,531 | | | | |

Table A.23. 2SLS regression of school enrolment of children

Robust standard errors in parentheses. Standard errors are clustered at age month of the oldest and household levels. *** p < 0.01, ** p < 0.05, * p < 0.1.