

Black, Sandra E.; Devereux, Paul J.; Lundborg, Petter; Majlesi, Kaveh

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## Poor Little Rich Kids? The Determinants of the Intergenerational Transmission of Wealth

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Department of Economics  
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Sandra E. Black  
Paul J. Devereux  
Petter Lundborg  
Kaveh Majlesi

August 2015



**LUND**  
UNIVERSITY

# **Poor Little Rich Kids?**

## **The Determinants of the Intergenerational Transmission of Wealth<sup>1</sup>**

by

Sandra E. Black  
Department of Economics  
University of Texas, Austin  
NHH, IZA and NBER  
sblack@austin.utexas.edu

Paul J. Devereux  
School of Economics and Geary Institute, University College Dublin  
CEPR and IZA  
devereux@ucd.ie

Petter Lundborg  
Lund University  
IZA and CED Lund

Kaveh Majlesi  
Department of Economics, Lund University  
Knut Wicksell Centre for Financial Studies

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### **Abstract**

Wealth is highly correlated between parents and their children; however, little is known about the extent to which these relationships are genetic or determined by environmental factors. We use administrative data on the net wealth of a large sample of Swedish adoptees merged with similar information for their biological and adoptive parents. Comparing the relationship between the wealth of adopted and biological parents and that of the adopted child, we find that, even prior to any inheritance, there is a substantial role for environment and a much smaller role for pre-birth factors. We also examine the role played by bequests and find that, when they are taken into account, the role of adoptive parental wealth becomes much stronger. We find very little evidence that education or earnings of parents or children are important drivers of the intergenerational wealth relationship between children and their adoptive parents. Our findings suggest that wealth transmission is not primarily because children from wealthier families are inherently more talented or more able but that, even in relatively egalitarian Sweden, wealth begets wealth.

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<sup>1</sup> The data used in this paper come from the Swedish Interdisciplinary Panel (SIP) administered at the Centre for Economic Demography, Lund University, Sweden.

## 1. Introduction

Wealth inequality has increased dramatically in recent decades. Indeed, a recent study found that, in the U.S., the median net worth of upper-income families doubled in a 30-year period, but declined for lower-income families.<sup>2</sup> This fact, in conjunction with the release of Thomas Piketty's Capital in the 21<sup>st</sup> Century that highlights the intergenerational transmission of wealth as a key determinant of the nature of society more generally, has brought renewed interest in understanding the determinants of the intergenerational correlation in wealth (Piketty 2014).

Given this, it is surprising how little we know about the nature of intergenerational correlations in wealth. While there are many studies about intergenerational transmission of education and income, much less is known about wealth, despite the fact that wealth is probably a better measure of economic success than income or education and is more easily transferred across generations.<sup>3</sup> Importantly, wealth is also much less equally distributed than education and income.

There have been a number of recent papers that have generated estimates of the intergenerational correlation of wealth. Charles and Hurst (2003) use U.S. data and find elasticities of about 0.37 for net wealth.<sup>4</sup> More recently, Boserup, Kopczuk, and Kreiner (2014) and Adermon, Lindahl, and Waldenström (2015) have used register data (from Denmark and Sweden, respectively) to estimate multi-generational models of wealth transmission and found strong positive intergenerational rank correlations.<sup>5</sup>

But why is wealth correlated across generations? Is it nature or nurture? There

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<sup>2</sup> <http://www.pewresearch.org/fact-tank/2014/12/17/wealth-gap-upper-middle-income/>

<sup>3</sup> See Black and Devereux (2011) for a survey of the literature on intergenerational mobility.

<sup>4</sup> See also Pfeffer and Killewald (2015) for a re-analysis using updated PSID data and Mulligan (1997) for earlier estimates of the intergenerational wealth elasticity in the U.S.

<sup>5</sup> Also noteworthy, Clark and Cummins (2014) use rare surnames and estate records to show strong transmission of wealth in England over many generations.

is very little work examining the underlying causes of these correlations.<sup>6</sup> One possible pathway is through biology-- genetic inheritance of skills, attitudes, and preferences that correlate with higher wealth in each generation.<sup>7</sup> Another pathway is environment--wealthier parents may invest more in their children's human capital, help their children get better jobs, provide funding for business start-ups, or give financial gifts. In this paper, we attempt to disentangle the role of nature versus nature in the intergenerational transmission of wealth.

The nature/nurture distinction is important as it distinguishes between fundamentally different reasons for positive intergenerational wealth correlations. The nature or genetic channel suggests that these correlations can arise because children from wealthy families are inherently more talented and would be wealthier than others even without the advantage of growing up with wealthy parents. The nurture or environmental channel instead suggests that parental wealth enables children to acquire greater wealth either directly through inheritance or indirectly through human capital investments or other environmental influences. The distinction is of great importance for our perspective on the intergenerational wealth correlation.<sup>8</sup>

To provide insight into this issue, we take advantage of a unique feature of the Swedish adoption system whereby we observe both the biological and adoptive parents of adopted children. We use administrative data on the net wealth of a large sample of

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<sup>6</sup> A related literature has examined the genetic contribution to different components of wealth by comparing fraternal and identical twins. Using this strategy, Cronqvist and Siegel (2015) argue that genetic differences explain a substantial fraction of the variation in savings propensities and wealth at retirement. Barnea, Cronqvist, and Siegel (2010) find that one-third of the variation in the share in equities in financial portfolios is explained by a genetic effect.

<sup>7</sup> Evidence on genetic effects in risk aversion and risk-taking behavior is found in Cesarini et al. (2010), Kuhnen and Chiao (2009), Dreber et al. (2009), and Black et al. (2015).

<sup>8</sup> It may also matter for policy. For example, a tax on parental wealth is likely to have less implication for intergenerational mobility if the intergenerational wealth correlation is predominantly due to nature rather than nurture. However, even if the intergenerational correlation was wholly due to nature, this does not imply that it could not be affected by policy.

adopted children born between 1950 and 1970 merged with similar information for their biological and adoptive parents--as well as corresponding data on own-birth children. We disentangle the role of nature versus nurture in the intergenerational transmission of wealth by looking at how the wealth of adoptive children is related to that of both their biological and adoptive parents. Adoption allows us to examine the effects of environmental factors in a situation where children have no genetic relationship with their (adoptive) parents.

We find that, even before any inheritance has occurred, wealth of adopted children is more closely related to the wealth of their adoptive parents than to that of their biological parents. This suggests that wealth transmission is primarily due to environmental factors rather than because children of wealthy parents are inherently more talented. We also examine the role played by bequests and find that, when they are taken into account, the role of adoptive parental wealth becomes much stronger.

The structure of the paper is as follows. In the next section, we discuss the institutional background both in terms of financial markets and the adoption process. In section 3, we outline the econometric methodology and, in section 4, we describe the data. Section 5 provides our estimates for the intergenerational transmission of wealth. In Section 6, we discuss a variety of robustness checks, including tests for non-random assignment of adoptees, the effects of varying the age of measurement of wealth, and different measures of net wealth. Section 7 then discusses our findings and concludes.

## **2. Institutional Background**

### **A. Wealth in Sweden**

Non-retirement wealth in Sweden is principally held in real assets--primarily housing--and financial wealth, including cash, stocks, and bonds. While we do not have

information on pension wealth, non-retirement wealth accounts for almost 84 percent of aggregate household financial wealth and a much higher proportion of total wealth.<sup>9</sup> However, it is important to understand the nature of the pension system due to its potential effect on savings.

Relative to countries such as the U.S., Sweden's pension system would be considered quite generous. Sweden has a mix of public and private pension schemes, and individuals are allocated to different pension systems depending on the public or private sector affiliation and year of birth of the individual. The longer one works, the higher the pension one receives. The retirement age is flexible and individuals can claim retirement benefits beginning at age 61.<sup>10</sup>

Because we study the individual wealth of children, it is important to understand whether there are incentives to transfer wealth holdings from one spouse to another. There do not appear to be any such incentives. In the event of a divorce, in the absence of a prenuptial agreement, all assets are split equally among spouses. For wealth tax purposes, the value of jointly owned assets was split evenly between the two tax filers. Thus, there were no incentives for husbands and wives to strategically allocate assets between themselves in order to reduce their wealth tax bill.

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<sup>9</sup> See Calvet, Campbell, and Sodini (2007). Also, stock market participation rates are higher in Sweden than in many other countries such as the United States (Guiso, Haliassos, and Jappelli, 2001).

<sup>10</sup> In 1999, when we measure wealth of parents, the public pension system almost entirely consisted of a national pension plan financed on a pay-as-you-go basis (an individual account system known as the Premium Pension System (PPS) was introduced in 1999). In addition, most people receive an occupational pension from their employer. According to the Swedish Pensions Agency, about 90% of employees receive some pension benefits from their employer as a condition of employment. On average, around 4.5% of the employee's salary is put into employer provided schemes (Thörnqvist and Vardardottir, 2014). Swedish residents also have tax incentives to invest in private pension savings that are only accessible after retirement. However, as mentioned earlier, individuals still hold a substantial fraction of their wealth in non-retirement wealth. There is also a guaranteed pension for those who have had little or no income from work, and the size of this guaranteed pension is based on how long the person has lived in Sweden. In 2000, the maximum guaranteed pension, which applies to those who have lived in Sweden for at least 40 years, is 2394 SEK per month (\$254) before taxes for those who are married, and 2928 SEK per month (\$311) for a single person. A tax rate of 30 percent is then applied.

## B. The Adoption System<sup>11</sup>

The adoptees we study were born between 1950 and 1970. During this period, private adoptions were illegal, so all adoptions went through the state. The state collected information on both the biological and adoptive parents; while it only required information on the biological mother, in many cases, social workers were also able to identify the biological fathers. While we do not observe how old the children were when they were adopted, about 80% of children were adopted in their first year of life.<sup>12</sup>

In order to adopt a child, a family had to satisfy certain requirements. The adoptive parents had to be married and be at least 25 years old, have appropriate housing, and be free of tuberculosis and sexually transmitted diseases. The adoptive father was required to have a steady income and the adoptive mother was expected to be able to stay home with the child for a certain period of time.<sup>13</sup> Overall, the adoption criteria meant that the adoptive parents were positively selected relative to the general population.

While matching of children to adoptive parents was at the discretion of the caseworkers, the evidence from that period suggests that social authorities were not able to systematically match babies to families based on family and child characteristics (see Lindquist, Sol, and Van Praag 2015 for more details).<sup>14</sup> However, we will examine

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<sup>11</sup> See Bjorklund, Lindahl, and Plug (2006) and Lindquist, Sol, and Van Praag (2015) for more details.

<sup>12</sup> Upon turning 18, an adopted child has the legal right to obtain information from public authorities about the identity of his or her biological parents (Socialstyrelsen 2014). However, according to Swedish law, there is no legal requirement for parents to inform adopted children that they are adopted (SOU 2009).

<sup>13</sup> Prior to 1974, there was no parental leave to care for adopted children. However, from 1955, mothers of biological children had a right to 3 months of paid leave (SOU 1954).

<sup>14</sup> While children could be adopted by relatives, in practice this was very rare. Nordlöf (2001) reports that of 1186 adoptions taking place in Stockholm for children born between 1960 and 1973, in only 4 cases was an adoptive parent related to the child.



this issue in more detail later.

### 3. Empirical Strategy

A large body of literature in economics has used data on adoptees to disentangle the relative contribution of genes and environment to economic behavior. These studies have typically used information on foreign-born adoptees, where the characteristics of the biological parents are unknown to the researcher. These studies have therefore not been able to compare the relative influence of biological and adoptive parents.<sup>15</sup>

However, a recent literature has taken advantage of the unique Swedish register data that identify both biological and adoptive parents. The first was the seminal study by Björklund, Lindahl, and Plug (2006), who studied the relative roles of nature versus nurture in the intergenerational transmission of educational attainment and earnings using cohorts born between 1962 and 1966. This was followed by papers using a similar strategy to study voting behavior (Cesarini, Johannesson, and Oskarsson, 2014), crime (Hjalmarsson and Lindquist, 2013), entrepreneurship (Lindquist, Sol, and Van Praag, 2015), health (Lindahl et al. 2015), and risk-taking in financial markets (Black et al. 2015). In general, these studies have found evidence that both characteristics of biological and adoptive parents are predictive of child outcomes.

Our main variable of interest is net wealth, which is constructed by subtracting total debts from total wealth. We transform the measure of net wealth in various ways in our empirical analyses. As our primary variable of interest, we construct within-cohort measures of parents' and children's rank within the wealth distribution. As

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<sup>15</sup> While working on this paper, we became aware of concurrent work by Fagereng, Mogstad, and Rønning (2015), who use Korean adoptees in Norway to determine the effect of environment on child wealth and asset allocation. The authors find a substantial role for environment. A key advantage of this work is that the assignment of children to families is arguably random. A key limitation, however, is that they do not observe characteristics of the biological family. We view this paper as a complement to our own work. More broadly, see Sacerdote (2010) for a survey of this literature.

discussed in more detail later, we base this choice on the fact that the relationship between child's rank and parent's rank is approximately linear. However, we also test the sensitivity of our conclusions to the choice of an inverse hyperbolic sine transformation of net wealth, as well as the untransformed value of net wealth (in levels).

Our main specification relates the rank of net wealth of an adoptee to the rank of net wealth of both his/her biological and adoptive parents. We estimate the following equation:

$$W_{ij} = \beta_0 + \beta_1 W_i + \beta_2 W_j + X\beta_3 + \epsilon_{ij} \quad (1)$$

where  $W$ , our main variable of interest, is the rank of net wealth,  $i$  indexes the biological family,  $j$  indexes the adoptive family, and  $X$  refers to the set of control variables. These include year-of-birth dummies for both parents and children and a dummy variable for the gender of the child. We measure child wealth at the individual level but measure parental wealth as the average of the mother's and father's wealth.<sup>16</sup>

For each child, we compute his/her rank in the distribution of child wealth for individuals born in the same year and so measured at the same age. Within an age cohort, ranks are normalized to lie between 0 and 100.<sup>17</sup> We use the child's rank in the entire distribution (of their cohort) throughout the analysis even when we are studying subgroups of children such as the sample of adoptees. We carry out the same exercise for parental wealth basing the cohort on the average cohort of the two parents. Because ranks of parents and children are uniformly distributed, the regression coefficients from equation (1) can be interpreted as rank correlations.

A key assumption of our empirical strategy is that adoptees are randomly

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<sup>16</sup> Charles, Hurst, and Killewald (2013) show that, in the U.S., wealthier individuals are likely to marry wealthier spouses.

<sup>17</sup> Ranks are calculated as  $[(i - 0.5)/N] * 100$  where  $i$  denotes individuals sorted by wealth, and  $i = 1, 2, \dots, N$ .

assigned to adoptive families at birth. If this assumption holds, the coefficients on the wealth of biological parents provide an estimate of the effect of pre-birth factors and the coefficient of adoptive parents provide an estimate of the effects of post-birth factors. The assumption will be violated if adoptees are systematically matched to adoptive parents that are similar to their biological parents. Previous studies using these adoption data have shown that, while children are clearly not assigned randomly to adoptive parents, the resultant biases are likely to be small. Following this literature, we will conduct a battery of robustness checks, where we provide evidence suggesting that any violations of the assumption do not invalidate our estimates.

#### 4. Data

We construct our database by merging a number of administrative registers. Our starting point is an administrative dataset containing information on all Swedish citizens born between 1932 and 1980. These data include information on educational attainment, county of residence, and other basic demographic information.<sup>18</sup> To this, we merge data from the Swedish multigenerational register, where we are able to identify Swedish-born adoptees by using information on both biological and adoptive parents of children.<sup>19</sup>

Our data on wealth come from the Swedish Wealth Data

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<sup>18</sup> We impute years of schooling based on the information on highest educational degree completed contained in the education register. We follow the coding of Holmlund, Lindahl, and Plug (2011) and impute years of schooling in the following way: 7 for (old) primary school, 9 for (new) compulsory schooling, 9.5 for (old) post-primary school (realskola), 11 for short high school, 12 for long high school, 14 for short university, 15.5 for long university, and 19 for a PhD university education. Since the education register does not distinguish between junior-secondary school (realskola) of different lengths (9 or 10 years), it is coded as 9.5 years. For similar reasons, long university is coded as 15.5 years of schooling.

<sup>19</sup> We know the identity of biological fathers for only about 50% of adoptees. While we cannot directly examine the sensitivity of our estimates to this issue because our parental variable is measured at the family level, previous studies that examined mother characteristics and behavior have found no evidence of bias due to missing fathers. See, for example, Björklund, Lindahl, and Plug (2006), Black et al. (2015), and Lindqvist et al. (2015).

(Förmögenhetsregistret). These data were collected by the government's statistical agency, Statistics Sweden, for tax purposes between 1999 and 2007, at which point the wealth tax was abolished.<sup>20</sup> For the years 1999 to 2006, the data include all financial assets held outside retirement accounts at the end of a tax year, December 31st, reported by a variety of different sources, including the Swedish Tax Agency, welfare agencies, and the private sector. Financial institutions provided information to the tax agency on their customers' security investments and dividends, interest paid or received, and deposits, including nontaxable securities and securities owned by investors, even for persons below the wealth tax threshold. Because the information is based on statements from financial institutions, it is likely to have very little measurement error, and because the entire population is observed, selection bias is not a problem.<sup>21</sup>

From the wealth register, we observe different categories of wealth. This includes the aggregate value of bank accounts, mutual funds, stocks, options, bonds, housing wealth, and capital endowment insurance as well as total financial assets and total assets.<sup>22</sup> The wealth register also contains data on total debt and net wealth.

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<sup>20</sup> During this time period, the wealth tax was paid on all the assets of the household, including real estate and financial securities, with the exception of private businesses and shares in small public businesses (Calvet, Campbell, and Sodini 2007). In 2000, the tax rate was 1.5 percent on net household wealth exceeding SEK 900,000. The Swedish krona traded at \$0.106 at the end of 2000, so this threshold corresponds to \$95,400. After 2000, the tax threshold was raised to SEK 1,500,000 for married couples and non-married cohabitating couples with common children and 1,000,000 for single taxpayers. In 2002 the threshold rose again, this time to SEK 2,000,000 for married couples and non-married cohabitating couples and 1,500,000 for single taxpayers. In 2005 the threshold rose once more but this time only for married couples and cohabitating couples, this time to SEK 3,000,000.

<sup>21</sup> In the case of foreign assets, individuals were required to report these themselves. Evidence suggests that unreported foreign assets likely represent a small fraction of total household assets. (Calvet, Campbell, and Sodini 2007)

<sup>22</sup> Small bank accounts were not reported by banks to the Swedish Tax Agency unless there was more than 100 SEK (about \$10) in interest during the year. However, Statistics Sweden estimates that 98% of the total money in bank accounts is included in the data.

Nonfinancial assets are collected from the property tax assessments and valuations are based on market prices.<sup>23</sup>

We measure income for our sample by using data from the Swedish Income Register. The register contains yearly income from 1968 onwards, and we use a measure of income that includes earnings from employed labor as well as self-employment income and taxable benefits.

We limit our analyses to children born 1950-1970 with all applicable parents alive in 1999 and for whom we have information on schooling, earnings, and wealth. In our analyses, we measure net wealth of the children in 2006 and net wealth of the parents in 1999. In order to avoid the issue of inheritances, we further restrict that at least one parent is alive in 2006 (for adoptees, we require that at least one adoptive parent be alive in 2006); however, we later test the sensitivity of our conclusions to this choice. The logic for restricting our sample to children born by 1970 and measuring their wealth in the latest possible year, 2006, is to avoid having very young people in the sample who have not yet had much opportunity to accumulate wealth. The average age of children in our sample is 44. This compares with an average age of 38 in Charles and Hurst (2003), 47 for the third generation in Adermon et al. (2015), and 34 for the second generation in Boserup et al. (2014). Later, we show that our estimates are not sensitive to the exact ages of the children at wealth measurement.

We have information on over 1.2 million children who are raised by their biological parents and 2519 adopted children for whom we have data available for both biological and adoptive mothers and fathers. Descriptive statistics for our sample are shown in Table 1. In the top panel, we show means for children, both biological and

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<sup>23</sup> Statistics Sweden adjusts tax-assessed property values using information on both tax assessments and actual sales prices of houses so the aggregate value of the housing stock in the data is consistent with sales prices (Adermon et al. 2014).

adoptive. In 2006, when their assets and education are measured, the average child age is 44 for biological children and 43 for adoptive children. On average, biological children have 0.4 of a year more education and hold slightly higher net wealth (621K SEK vs. 591K SEK).

In the second panel, we show means for biological parents, both parents who raised their own biological children and parents who gave their children up for adoption. The two types of parents are quite different in their characteristics, with biological parents of adoptees being much less wealthy and having fewer years of schooling.

The bottom panel of Table 1 shows descriptive statistics for adoptive parents. For adopted children, adoptive parents are, on average, older, wealthier, and better educated than the child's biological parents. Adoptive parents also appear positively selected when we compare them to biological parents who raise their own children, although the differences here are much smaller.<sup>24</sup>

## 5. Results

When considering the intergenerational correlation in wealth, the literature is agnostic as to the appropriate functional form. Research in the area has used a variety of transformations of net wealth, including levels, logs, the inverse hyperbolic sine transformation, and within cohort ranks. When we examine the data, it is clear that the

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<sup>24</sup> Appendix Table 1 provides a breakdown of the proportions of sample members who have positive, zero, and negative net wealth respectively. In the sample of own-birth children, almost 1% have zero net wealth, and 26% have negative wealth. For adoptive children, the percentages are 1% and 32%. As discussed by Boserup et al. (2015), standard life-cycle theory would predict negative wealth for young persons with increasing earnings profiles. Unsurprisingly, the proportions with zero and negative wealth are lower for parents, both because they are older and because we are averaging wealth across the father and the mother. Among parents of own-birth children, 11.1% have negative wealth and 1% have zero net wealth. The percentages with negative wealth are 5% for adoptive parents and 31% for biological parents of adoptees. This provides further evidence that biological parents of adoptees are negatively selected.

within-cohort rank specification best fits the linear model; as a result, we use that as our preferred specification. However, in later analyses, we will show that our conclusions are robust to the choice of the measure of net wealth.

Figure 1a plots the relationship between the within-cohort rank of net wealth of parents and children for the large own-birth sample using a local linear kernel regression with an epanechnikov kernel and rule-of-thumb bandwidth.<sup>25</sup> Importantly, we see that this relationship is approximately linear from around the 5<sup>th</sup> percentile to the 95<sup>th</sup> percentile. Consistent with the Swedish findings of Adermon et al. (2015), the slope is negative at the very bottom of the distribution and more steeply positive at the top. The declining slope at the bottom is driven by parents with large negative wealth. The increase in slope at the top is consistent with general findings of greater persistence in economic status at the very top of the distribution (Björklund et al. 2012). Figure 1b shows the equivalent picture when we drop the parents in the top and bottom 5% of their within-cohort distribution, and the linearity of the relationship becomes more pronounced.<sup>26</sup>

Among adopted children, Figures 2a and 2b plot the within-cohort rank relationship between children and biological and adoptive parents respectively. Here, we see similar patterns to the full sample. However, confidence intervals become much wider at the tails, and this is more pronounced at the top of the distribution among biological parents and at the bottom of the distribution among adoptive parents. This highlights the fact that biological parents are primarily negatively selected in terms of

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<sup>25</sup> Adermon et al. (2015) also use this approach. An alternative, used by Boserup et al. (2014), is to plot average child rank against parental wealth percentile. The local linear kernel regression is more efficient and this is important given our sample of adoptees is not very large.

<sup>26</sup> In this case, we rank all individuals within a given cohort (for parents, we calculate the cohort as the rounded average of the father and mother) and trim the top and bottom 5%. As a result, adoptive parents, biological parents, and own-birth parents in the same cohort are all ranked within the same distribution.

net wealth while adoptive parents are positively selected. When we trim the top and bottom 5% of the data, the relationship again becomes much more linear. (Figures 3a and 3b).

In Table 2, we report the regression results when we estimate equation (1) on the sample of own-birth children (Columns 1 and 3) and adoptees (Columns 2 and 4). As noted earlier, we include cohort dummies for parents and children in all specifications.<sup>27</sup>

Columns 1 and 2 present the rank-rank coefficient for own-birth and adopted children, respectively. In the case of adopted children we control for the within-cohort rank of the net wealth of biological parents as well as that of adoptive parents. Among own-birth children (Column 1), the rank-rank coefficient is approximately 0.35. This implies that a one percentile increase in the position of parents in the wealth distribution is associated with just over one third of a percentile increase in the average position of their children. Among adoptees (Column 2), we find that child's wealth is predominantly associated with that of adoptive parents and has a much weaker relationship with biological parents' wealth. The rank coefficient for biological parent wealth is 0.11 but that for adoptive parent wealth is 0.27.<sup>28</sup>

We saw in Figures 1 and 2 that the rank-rank relationship is approximately linear except in the tails of the parental wealth distribution -- for ranks up to the 5th percentile and in the very top of the distribution. Therefore, in Columns 3 and 4, we drop cases with parental wealth in the top or bottom 5 percentiles of the within-cohort parental wealth distribution. This is particularly important in the adoptive sample, as

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<sup>27</sup> Given wealth is measured in the same year for all parents and wealth is measured in the same year for all children, these also serve as age dummies. The estimates without these dummies are quite similar. This is what we would expect for the rank transformation as the ranks are computed by cohort.

<sup>28</sup> There are 460 own-birth children of adoptive parents in our data. The estimated effect of rank parental wealth on their wealth rank is 0.36, which is very similar to that for the full sample of own-birth children.



biological parents are much poorer than adoptive parents. Not surprisingly, given the figures earlier, these exclusions do affect our estimates, with an increase in the effect of biological wealth and a decrease in the effect of adoptive wealth. Still, however, the adoptive coefficient is substantially larger than the biological one. The relatively weak relationship between biological parental wealth and child wealth is interesting as it suggests that most of the reason for the intergenerational transmission of wealth is not due to the fact that children from wealthier families are inherently more talented. Instead, it appears that, even in a relatively egalitarian society like Sweden, wealth begets wealth.

We next consider whether these relationships are the same for sons and daughters. We do not have a strong prior in terms of whether adoptive or biological relationships should be stronger for boys or girls. In Table 3, we report the estimates for our preferred specification where we exclude children whose biological or adoptive parents have net wealth in the bottom or top 5% of the rank distribution. Columns 1 and 2 present the results by child gender. While the biological coefficient is larger for boys than for girls, the difference is not statistically significant. The adoptive coefficient is similar for both genders, suggesting there is not much evidence for gender differences in the nature/nurture split.

Finally, we consider the potential role of inheritances when estimating intergenerational correlations in wealth.<sup>29</sup> In Sweden, as in the United States, when a spouse dies their assets automatically transfer to the surviving spouse. Because we have restricted the sample so that at least one parent is alive when child wealth is measured in 2006, we are unlikely to have captured bequests. To test the potential role of inheritances, we compare two extreme cases--in one, at least one parent was alive in

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<sup>29</sup> Piketty and Zucman (2014, 2015) show that inheritance can have important effects on the distribution of wealth.

1999 but both parents are dead by 2006, suggesting that the child is likely to have received inheritances in the interim, and in the other case, both parents are alive in both periods, ruling out the possibility of inheritance. Column 3 of Table 3 presents the first scenario; to estimate the potential effect of inheritances, we add a dummy variable for whether both parents are deceased in 2006 plus an interaction of this dummy variable with adoptive parental wealth.<sup>30</sup> The estimates are in column (3) of Table 3. While we have added only about 100 extra adoptive families to the sample, we still find a statistically significant interaction effect of 0.42. This suggests that the rank correlation with adoptive parent wealth increases from 0.23 to 0.65 once inheritances are included.<sup>31</sup> This large effect is consistent with the findings of Adermon et al. (2015) who use wealth and inheritances data and find that inheritance appears to be the most important component of the intergenerational wealth elasticity in Sweden.

At the other extreme, we rule out the possibility that the child received an inheritance by restricting the sample to cases with both adoptive parents alive in 2006 (Column 4). While this reduces the sample size considerably, the estimates are largely unchanged from the baseline in Table 2 column 4. This is consistent with our expectation that bequests to children occur after both parents die.

## 6. Robustness Checks

### *Random Assignment of Adoptees*

As noted earlier, our identification strategy relies on the random assignment of adoptees. However, although adoptees are not randomly assigned to parents, we can

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<sup>30</sup> We assume that biological parents of adoptive children will not have bequest motives for the children they gave up.

<sup>31</sup> There was an inheritance tax in Sweden until December 2004 when it was abolished. When it was in effect, heirs paid a progressive tax rate of between 10% and 30% on inheritances above a 70,000 SEK deductible (about \$8000). This would tend to reduce the inheritance effect that we find.

evaluate how this non-random assignment might be affecting our estimates. The primary concern is that children may have been assigned to adoptive parents in such a way that there are correlations between net wealth of adoptive (biological) parents and unobserved characteristics of the biological (adoptive) parents that are correlated with child wealth. While earlier work using similar identification strategies and data suggest that this is unlikely to be a serious problem, we follow the literature in conducting a number of robustness checks to verify this.

If there are correlations between the wealth of adoptive parents and unobserved characteristics of the biological parents that are correlated with child wealth, the coefficients on wealth of adoptive parents may be sensitive to whether or not the wealth of biological parents is included in the regression. The results when we do this are presented in Table 4.

Column 1 of Table 4 shows estimates with just the wealth of the adoptive parents included.<sup>32</sup> In column 2, we add wealth of the biological parents, which is the specification previously reported in column 4 of Table 2. The coefficient on adoptive parent wealth changes very little when we include biological parent wealth, suggesting that the two variables are not highly correlated.<sup>33</sup>

As another check for omitted variable bias, we next include a number of other controls for characteristics of the biological parents; these include education and labor income and are included separately for mothers and fathers.<sup>34</sup> Column 3 of Table 4

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<sup>32</sup> All specifications include cohort dummies for parents and children.

<sup>33</sup> Note that a major advantage of having information on biological parents is that we can control for biological parental wealth in our specifications. To the extent that the adoption authorities attempt to match children to parents who have similar characteristics to the biological parents, controlling for biological parental wealth will reduce any bias in the coefficient on adoptive parental wealth resulting from non-random assignment.

<sup>34</sup> Our measure of biological parent earnings is calculated separately for mothers and fathers and is the log of average income between the years 1980 and 1999. In the very few cases where parental labor income is zero in all years, we set the log to zero.

includes wealth of adoptive parents and adds the further controls as proxies for general unobserved characteristics of biological parents. Comparing the coefficients on adoptive parent wealth in column 3 to column 1, again the difference is very small. Finally, in column 4, we include both biological parents' wealth as well as controls for their schooling and income. The resulting estimates are almost identical to those in column 3. Overall, it appears that our adoptive estimates are unlikely to be significantly biased by non-random assignment.

Columns 5 to 8 of Table 4 carry out the analogous exercise for wealth of biological parents. In column 5, we include only the wealth of biological parents and then systematically add controls for characteristics of adoptive parents. Column 6 includes controls for the wealth of adoptive parents, Column 7 includes controls for education and income (again entered separately for mothers and fathers), and Column 8 includes both sets of controls. While the coefficients on wealth of biological parents decrease somewhat in columns 6-8 compared to column 5, the differences are not very large. This suggests that non-random assignment of adoptees is unlikely to be a problem and, if anything, will lead to an overstatement of the role of biological parents relative to that of adoptive parents.

#### *Ages at Measurement of Wealth*

While we chose to measure wealth when the children were at their oldest (in 2006) to avoid them being too young to have accumulated wealth and when the parents were at their youngest (in 1999) in order to avoid issues of retirement, we next test the sensitivity of our conclusions to these choices. There is a clear life-cycle pattern to

wealth accumulation.<sup>35</sup> Figure 4a plots average net wealth by age for the full sample of parents with children born between 1950 and 1970. We see that average net wealth increases between age 50 and 60 and then is remarkably stable from the late 50s to the mid-80s. This is reassuring as it implies that the fact that many of our adoptive parents are quite old (the average age of adoptive parents is 68.6 in 1999) is unlikely to make their wealth levels unrepresentative. If anything, it may be that the biological parents of our adoptive children are a little young at measurement (average age is 59.6 in 1999); we show later that our results are robust to measuring their wealth in 2006 when their average age is 65.5.

By measuring child wealth in 2006, we are measuring it as late as possible in our data and all children are aged at least 36. However, there is still the concern that, because it is relatively early in the career for many of these children, our measure of wealth may not be representative of their wealth at later ages. Figure 4b plots child wealth by age. It is clear that all of our children (aged 36-56) are at ages at which average wealth is still increasing. However, the fact that we are not measuring child wealth at its maximum does not imply that our nature/nurture estimates are biased, as the relative importance of these factors may not change much over these ages.<sup>36</sup>

In Table 5, we investigate this issue by allowing for differential effects depending on the age at which wealth is measured. For children, we create a dummy equal to 1 if they are born between 1961 and 1970 (and so aged between 36 and 45 at wealth measurement and we interact this with wealth of both types of parents. We include these interactions in Column 1; in this specification, the main effects can be

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<sup>35</sup> Using Norwegian data, Fagereng, Gottlieb, and Guiso (2013) document life cycle patterns in stock market participation and portfolio allocation.

<sup>36</sup> This may be particularly true because we use within-cohort rank as our measure of net wealth. Nybom and Stuhler (2015) show that, in the case of income, the intergenerational rank correlation is much more robust to age at measurement than is the intergenerational elasticity.

interpreted as the effects of parental wealth for children aged between 46 and 56 at measurement. We see that the interaction effects are statistically insignificant and the main effects are similar to those in Table 2 Column 4.<sup>37</sup> This suggests that our estimates are not sensitive to the age of child at wealth measurement.

In Column 2 of Table 5, we similarly test whether the coefficient estimates depend on parental age at measurement. We define an older group of parents who are aged 65 or older at measurement and we interact this with parental wealth. The main effects can then be interpreted as the effects of parental wealth for the relatively young parents. Once again the interaction terms are statistically insignificant and the main effects are very similar to earlier estimates. In Column 3, we include interactions with the age dummies for both parents and children and once again find insignificant interaction terms. It appears that the relative contribution of nature and nurture is largely invariant to the exact age at measurement of wealth of parents and children in our sample.

Another potential issue is that biological parents are on average 9 years younger than the adoptive parents in 1999 (average age of 59.6 versus 68.6). Given that there are life-cycle patterns in wealth-holding, our conclusions may be sensitive to this difference. To address this, we measure the wealth of adoptive parents in 1999 and biological parents in 2006, thus largely eliminating the age gap at measurement. Column 4 of Table 5 reports these estimates. Once again, we find that the estimates are invariant to the age of measurement – the estimates in Column 4 of Table 5 are similar to our main specification in Table 2 Column 4.

#### *Different Transformations of Net Wealth*

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<sup>37</sup> We have also tried interactions using a continuous child age variable and found the interactions to be small and statistically insignificant.

Thus far, we have used the within-cohort rank as our measure of net wealth—from our own analysis, it is clear that this transformation fits the linear model the best. However, we next test the sensitivity of our conclusions to this choice. In addition to within-cohort rank, we consider the inverse hyperbolic sine transformation as well as the level of net wealth.<sup>38</sup>

Charles and Hurst (2003) use a log transformation for both parent and child wealth. However, this requires excluding all cases in which either parent or child has zero or negative net wealth and many individuals have non-positive net wealth. To avoid using a selected sample, we use the inverse hyperbolic sine transformation (IHS) rather than logs.<sup>39</sup> The IHS transformation of wealth,  $W$ , is  $w = \log(W + \sqrt{W^2 + 1})$  and behaves as  $\log(W)$  for positive values.<sup>40</sup>

Appendix Table 2 presents the results when we estimate equation 1 using these alternative measures of net wealth as our variables of interest. The IHS estimates for own-birth children (Columns 1 and 3) suggest an intergenerational elasticity of about 0.28—the results are relatively constant whether the data are trimmed or not. Among adoptees, we find similar patterns (Columns 2 and 4), with coefficients of .10 on biological parents' wealth and .25 on adoptive parents' wealth, and these relative patterns change little when we trim the data.

Columns 5-8 show the relationship between parental and child net wealth when wealth is not transformed and is simply reported in levels. The levels estimate among own-birth children is about 0.4 in the full sample but jumps to 0.6 when we exclude

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<sup>38</sup> Graphs of the relationship between parents and children's net wealth using these alternative transformations do not, in fact, look linear; as a result, we chose to use the within-cohort rank as our preferred specification. These figures are available from the authors upon request.

<sup>39</sup> The IHS is advocated by Pence (2006) as a superior alternative to using logs when studying wealth data.

<sup>40</sup> We have verified in our data that the relationship between parent and child net wealth using the IHS is exactly the same as that using logs once all negative and zero values have been excluded.

wealth levels in the bottom and top 5% of the distribution of ranks (Columns 5 and 7). This large change reflects the underlying non-linearities in the data. Finally, when we consider adoptees, the adoptive parent coefficient is 0.41 and the biological coefficient is 0.07 in the full sample; once we trim the data, the coefficient on biological parental wealth almost triples to 0.19 compared with 0.44 for adoptive parental wealth. We place little credence on the untrimmed estimates for the levels specification, however, given the sensitivity to outliers. Overall, our conclusions of the relative importance of adoptive parent's wealth relative to that of biological parents are robust to the choice of specification for net wealth.

#### *Other Robustness Checks*

We report further robustness checks in Appendix Table 3. Column 1 presents the baseline results from Table 2 Column 4 for comparison. In Column 2, we consider whether our conclusions are sensitive to correlations between wealth and residence. It may be that the wealth of parents and children are correlated because both live in an area that has high wealth levels -- for example, they may both live in an area with high property values. To examine this, in Column 2, we add controls for county of residence of both parents and children in 2000.<sup>41</sup> This has no effect on the estimates.

We have also thus far assumed that the effects of biological and adoptive parents are independent of each other. However, this may be an oversimplification if there are nature/nurture interactions, one building on the other.<sup>42</sup> We present the results when we allow for an interaction between biological and adoptive parents in column (3)

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<sup>41</sup> Sweden is divided into 20 regional county councils. Their main responsibilities are to provide and organize health care and public transportation.

<sup>42</sup> There are mixed findings in the literature about these types of interactions – Bjorklund, Lindahl, and Plug (2006) finds evidence of these interactions for mothers' education and fathers' earnings but Lindquist, Sol, van Praag (2015) find no evidence for these interactions when studying entrepreneurship and Black et al. (2015) find no evidence for them when studying risky investment behavior.



of Appendix Table 3. The interaction term is positive but statistically insignificant and so provides no evidence for a nature/nurture interaction.

Finally, while our wealth data are high quality and unlikely to suffer from significant measurement error, there could be transitory shocks to wealth that lead our estimates based on single years of wealth data to be misleading. Therefore, in column (4) of Appendix Table 3, we measure child wealth as the average in 2004-06 and parental wealth as the average over 1999-2001. We find that the averaging makes no appreciable difference to the estimates.

## **7. Conclusions**

There is an extensive body of research documenting a correlation in wealth across generations, with limited understanding of the underlying causes of this relationship. Taking advantage of unique data from Sweden that link adopted children to both their biological and adoptive parents, we are able to disentangle the role of nature versus nurture in the intergenerational transmission of wealth.

We find a substantial role for environmental influences with a much smaller role for biological factors, suggesting that wealth transmission is not primarily because children from wealthier families are inherently more talented or more able. Instead, it suggests that pre-birth endowment is a relatively small factor in this intergenerational relationship. We also find that when bequests are taken into account the role of adoptive parental wealth becomes much stronger. Importantly, our conclusions are robust to a variety of specification and robustness checks.

While we have established the relative role of nature versus nurture, the exact mechanisms of wealth transmission are more difficult to ascertain. Wealthier parents tend to be better educated and earn higher incomes, and these factors could lead to the

increased wealth of their children, through, for example, teaching them about investment opportunities or providing the right opportunities. However, when we investigate this, we find little evidence that this is the case.<sup>43</sup> It may also be that wealthy parents invest more in their child's education and career, which could then lead to higher child wealth accumulation. When we examine whether this is the case, however, we find little evidence for child education or income as mechanisms.<sup>44</sup> So, the pathway through which adoptive parental wealth affects child wealth does not appear to be primarily parental schooling and income or child human capital accumulation and greater labor earnings.<sup>45</sup> Taken together, our findings suggest potential roles for intergenerational transmission of behaviors or preferences (children of wealthier parents may choose to save more or invest in assets that have higher returns) or for financial gifts from parents to children. Unfortunately, we do not have information on savings behavior over the life-cycle or on financial gifts so this evidence is only suggestive.

It is clear from our results that pre-birth endowments do not drive the intergenerational correlations in wealth we observe; however, more work is required to determine the exact mechanisms through which wealthy parents create wealthy children.

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<sup>43</sup> To investigate whether this can explain the patterns we observe, in Appendix Table 4, we have added controls for adoptive parents' education and income, including them separately for fathers and mothers. This has negligible effects on the coefficients on parental wealth, suggesting that adoptive parental wealth has a direct effect on child wealth that does not come through other parental characteristics.

<sup>44</sup> We have examined the effects of parental wealth on child educational attainment and found the effect of adoptive parental wealth to be positive but small. Likewise, the effects of adoptive parental wealth on child labor earnings is modest and smaller than that of biological parents. Indeed, the effect of adoptive parental wealth on child wealth falls very little even when child education and labor earnings are introduced as additional controls (columns 3 and 4 of Appendix Table 4).

<sup>45</sup> This is perhaps surprising as Charles and Hurst (2003) find, in the U.S., that parental and child income can account for about half the intergenerational wealth elasticity.

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Figure 1a: Within-Cohort Wealth Rank Relationship between Parents and Own-birth Children

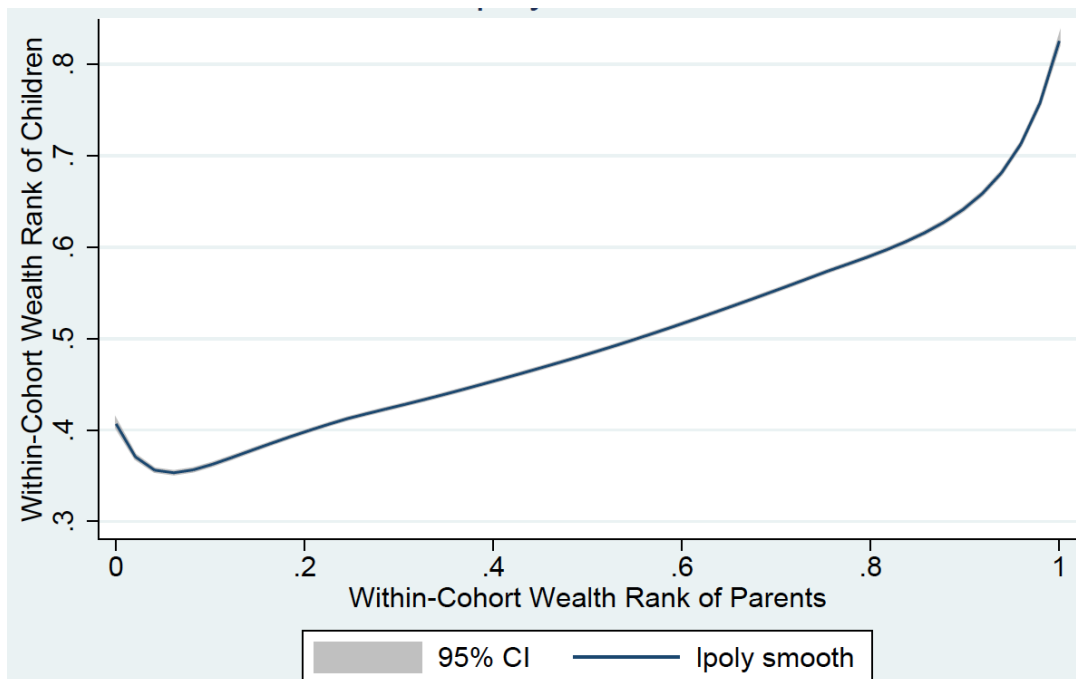


Figure 1b: Within-Cohort Wealth Rank Relationship between Parents and Own-birth Children  
Parents in the top and bottom 5% of the within-cohort wealth distribution are dropped

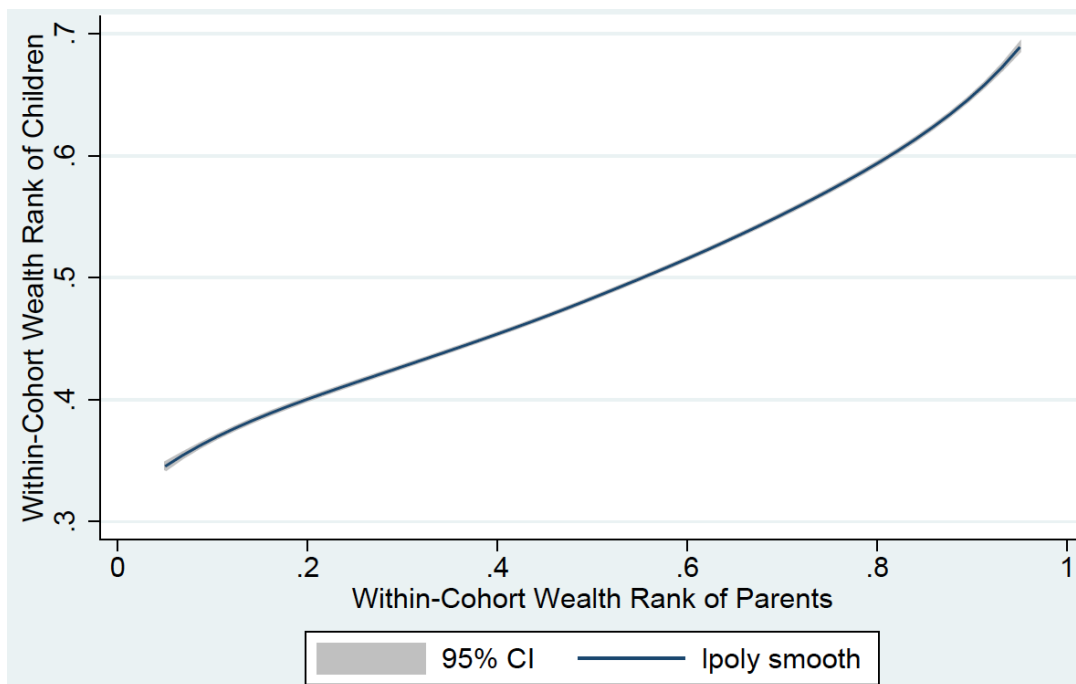


Figure 2a: Within-Cohort Wealth Rank Relationship between Adopted Children and Their Biological Parents

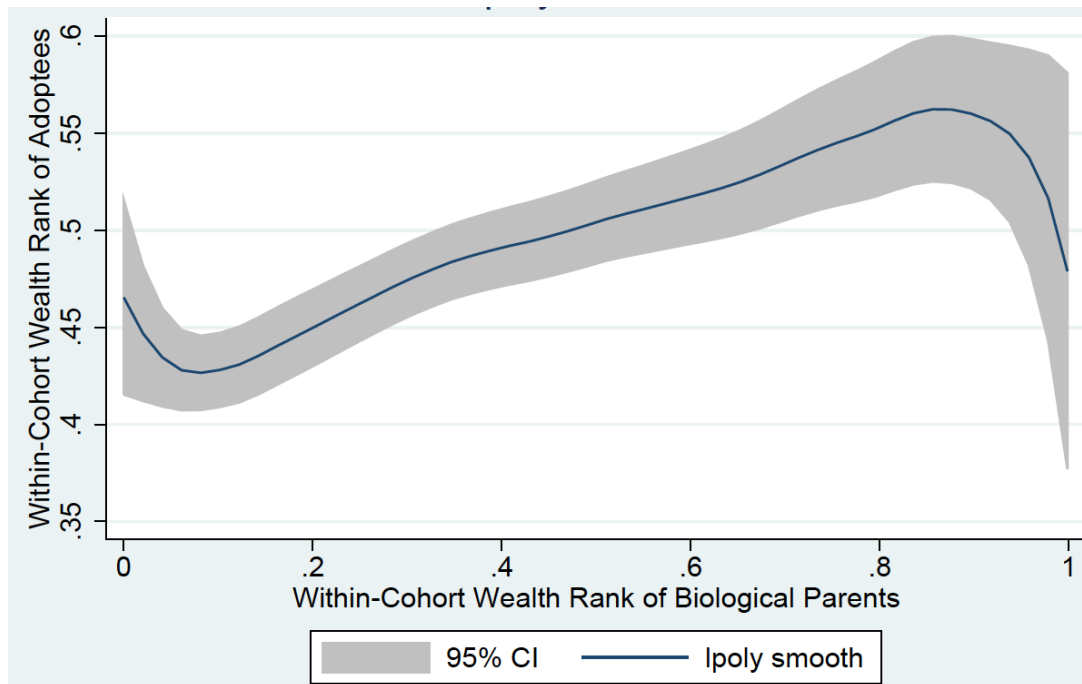


Figure 2b: Within-Cohort Wealth Rank Relationship between Parents and Children and Their Adoptive Parents

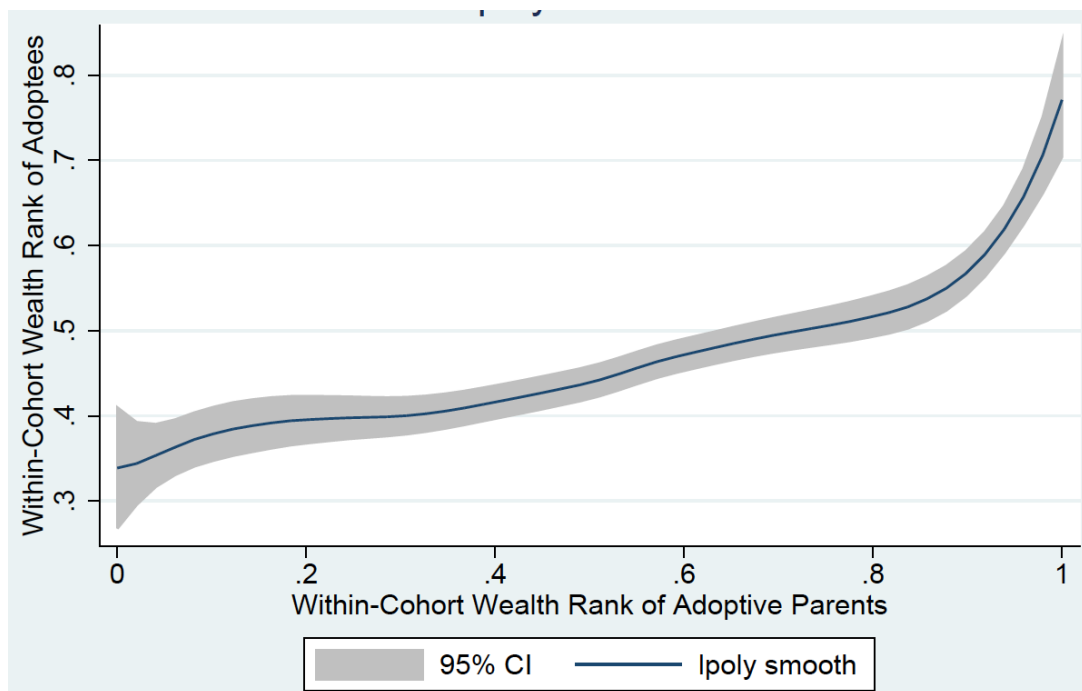


Figure 3a: Within-Cohort Wealth Rank Relationship between Adopted Children and Their Biological Parents- Parents in the top and bottom 5% of the within-cohort wealth distribution are dropped

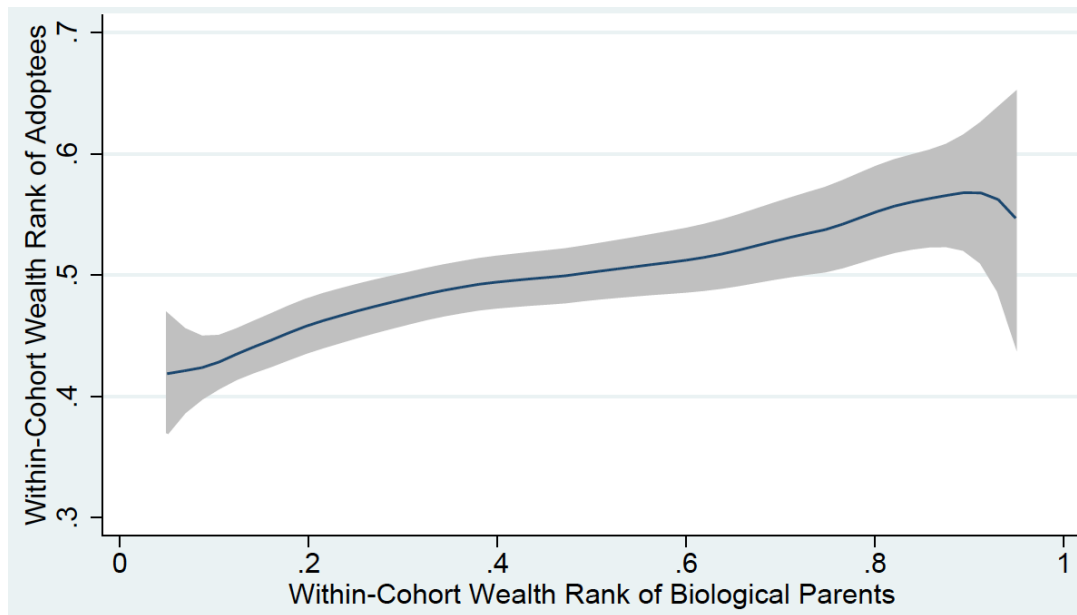


Figure 3b: Within-Cohort Wealth Rank Relationship between Parents and Children and Their Adoptive Parents- Parents in the top and bottom 5% of the within-cohort wealth distribution are dropped

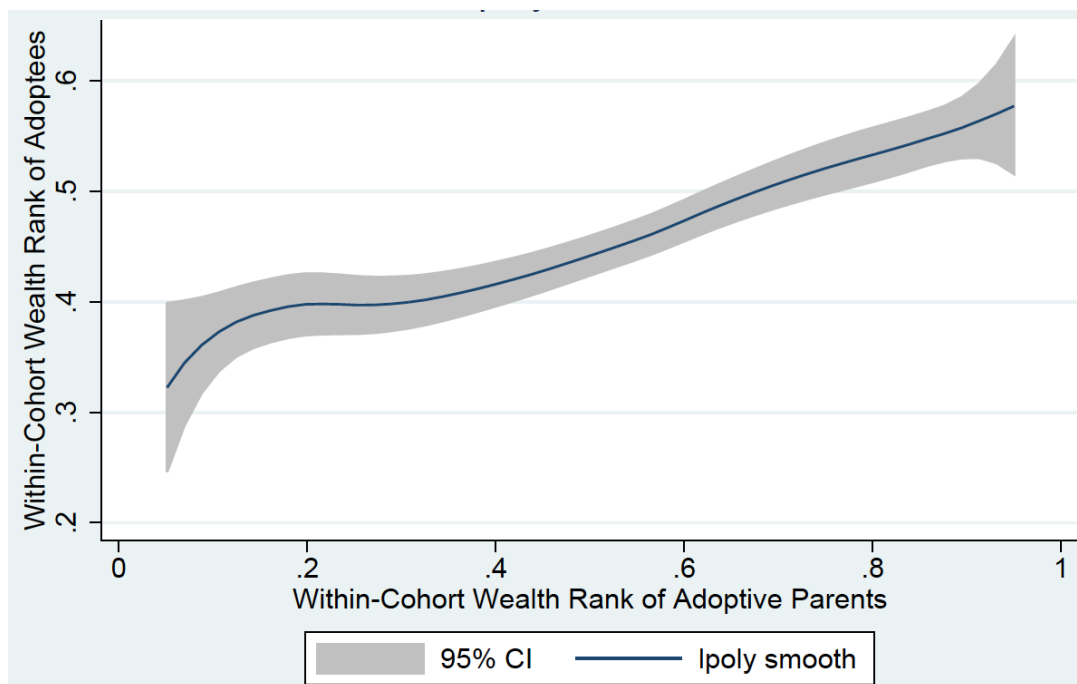
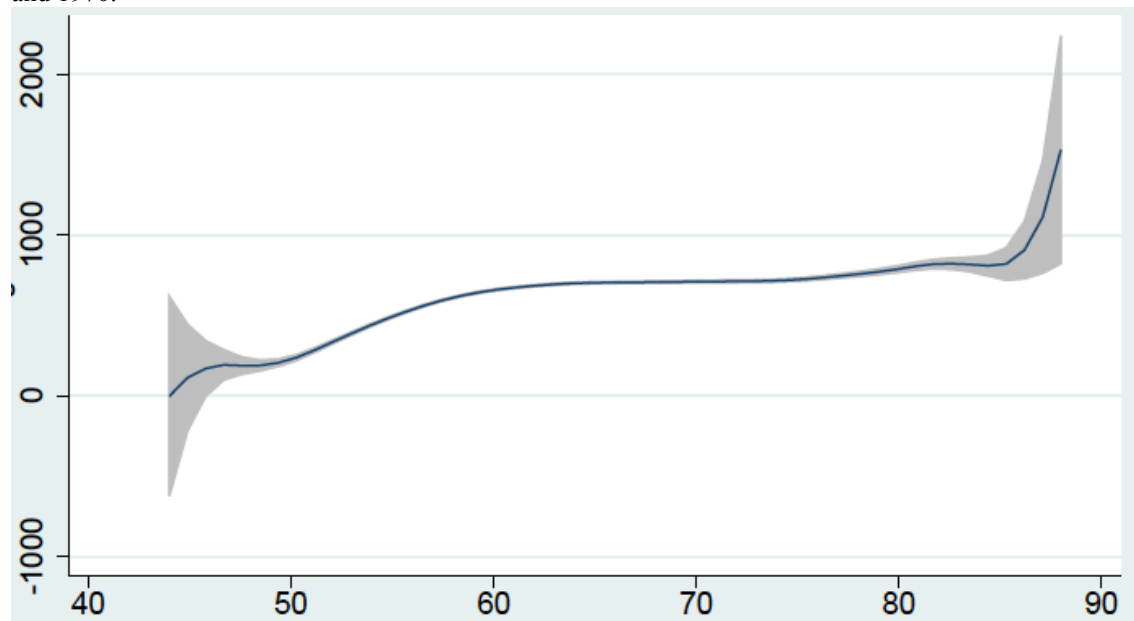


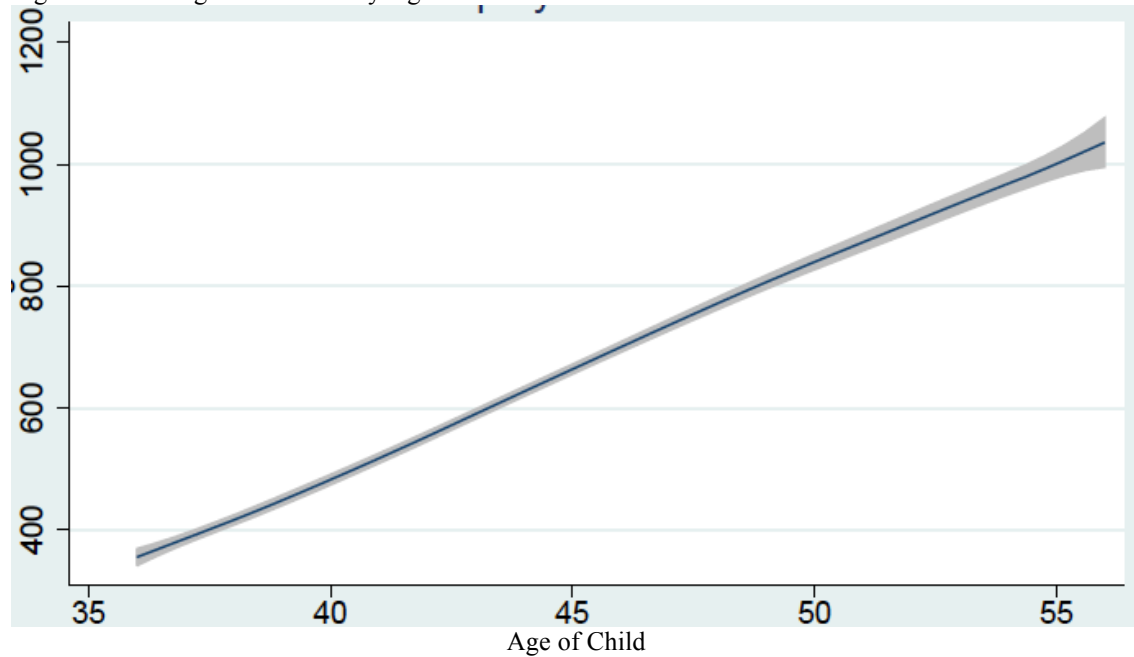


Figure 4a: Average Net Wealth by Age for the Full Sample of Parents with Children Born between 1950 and 1970.



Notes: Parental Net Wealth (in 1000 SEK). Line represents local linear approximation and shading represents the 95% confidence interval.

Figure 4b: Average Net Wealth by Age for Children Born between 1950 and 1970.



Notes: Child Net Wealth (in 1000 SEK). Line represents local linear approximation and shading represents the 95% confidence interval.

Table 1: Summary Statistics

	Own-birth children		Adopted children	
	Mean	SD	Mean	SD
	Children			
Net wealth rank	0.50	0.29	0.48	0.30
Net wealth*	620,757	2,893,464	591,463	1,597,692
Age in 2006	43.82	5.59	43.17	4.72
Years of schooling	12.37	2.30	11.98	2.12
Female	0.51	0.50	0.53	0.50
Observations	1,200,835		2,519	
	Biological parents			
Average net wealth ranking	0.50	0.29	0.34	0.27
Average net wealth*	640,802	2,009,063	243,999	667,880
Average age in 1999	63.94	7.43	59.58	6.62
Average years of schooling	10.13	2.62	9.65	2.08
	Adoptive parents			
Average net wealth ranking			0.55	0.28
Average net wealth*			826,294	2,233,660
Average age in 1999			68.63	6.49
Average years of schooling			10.51	2.82

Notes: \* Monetary values are reported in Swedish Krona on December 31, 2000. At the time, the exchange rate was 1 USD = 9.42 SEK. Parental wealth is calculated as combined wealth of the mother and father divided by two.

Table 2: Intergenerational Relationships  
Dependent Variable: Child Rank in Within-Cohort Wealth Distribution

	Full Sample		Trim Bottom/Top 5%	
	Biological Children	Adoptees	Biological Children	Adoptees
Rank Biological Parental Wealth	0.348 (0.001)***	0.113 (0.022)***	0.331 (0.001)***	0.132 (0.026)***
Rank Adoptive Parental Wealth		0.272 (0.021)***		0.229 (0.027)***
Observations	1,200,835	2,519	1,080,842	1,971
R squared	0.148	0.130	0.120	0.116

Notes: All specifications include cohort dummies for parents and children. Parental wealth is measured in 1999 and child wealth is measured in 2006. All parents are alive in 1999 and at least one of the (adoptive) parents of (adopted) biological children is alive in 2006. Parental wealth is calculated as combined wealth of the mother and father divided by two. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors clustered by adoptive family.

Table 3: Heterogeneous Effects  
Dependent Variable: Child Rank in Within-Cohort Wealth Distribution

VARIABLES	(1) Males	(2) Females	(3) Inheritance	(4) Both Adoptive Spouses Alive in 2006
Rank Biological Parental Wealth	0.162 (0.038)***	0.117 (0.037)***	0.124 (0.026)***	0.131 (0.030)***
Rank Adoptive Parental Wealth	0.222 (0.037)***	0.243 (0.040)***	0.231 (0.027)***	0.241 (0.031)***
Both adoptive parents died by 2006* Rank Adoptive Parental Wealth			0.417 (0.096)***	
Observations	1,037	934	2,046	1,531
R-squared	0.159	0.162	0.133	0.132

Notes: The top and bottom 5 percent of parental wealth distribution have been dropped. All specifications include cohort dummies for parents and children. Parental wealth is measured in 1999 and child wealth is measured in 2006. All parents are alive in 1999. In Columns (1) and (2) at least one adoptive parent is alive in 2006. Parental wealth is calculated as combined wealth of the mother and father divided by two. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors clustered by adoptive family.

Table 4: Addressing The Non-random Assignment of Adoptees  
Dependent Variable: Child Rank in Within-Cohort Wealth Distribution

VARIABLES	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Rank Biological Parental Wealth		0.132 (0.026)***		0.115 (0.027)***	0.159 (0.027)***	0.132 (0.026)***	0.153 (0.027)***	0.138 (0.026)***
Rank Adoptive Parental Wealth	0.239 (0.026)***	0.229 (0.027)***	0.231 (0.027)***	0.224 (0.027)***		0.229 (0.027)***		0.251 (0.027)***
Observations	1,971	1,971	1,971	1,971	1,971	1,971	1,971	1,971
R-squared	0.090	0.116	0.097	0.119	0.045	0.116	0.049	0.091
Biological Parents' Chars	NO	NO	YES	YES	NO	NO	NO	NO
Adoptive Parents' Chars	NO	NO	NO	NO	NO	NO	YES	YES

Notes: The top and bottom 5 percent of parental wealth distribution have been dropped. All specifications include cohort dummies for parents and children. Parental wealth is measured in 1999 and child wealth is measured in 2006. All parents are alive in 1999 and at least one adoptive parent is alive in 2006. Parental wealth is calculated as combined wealth of the mother and father divided by two. Parental Characteristics include Schooling and Income. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors clustered by adoptive family.

Table 5: Measuring Wealth at Different Ages  
Dependent Variable: Child Rank in Within-Cohort Wealth Distribution

VARIABLES	(1)	(2)	(3)	(4) Biological Parents' wealth measured in 2006
Rank Biological Parental Wealth	0.125 (0.048)***	0.128 (0.030)***	0.112 (0.059)*	0.108 (0.028)***
Rank Adoptive Parental Wealth	0.258 (0.048)***	0.200 (0.048)***	0.232 (0.075)***	0.224 (0.030)***
Rank Biological Parent Wealth * Child Aged 36-45	0.010 (0.058)		0.020 (0.063)	
Rank Adoptive Parent Wealth * Child Aged 36-45	-0.042 (0.058)		-0.034 (0.062)	
Rank Bio Parent Wealth*Bio Parent Aged 65+		0.019 (0.062)	0.028 (0.068)	
Rank Ad Parent Wealth * Ad Parent Aged 65+		0.039 (0.057)	0.027 (0.061)	
Observations	1,971	1,971	1,971	1,496
R-squared	0.116	0.116	0.116	0.148

Notes: The top and bottom 5 percent of the parental wealth distribution have been dropped. All specifications include cohort dummies for parents and children. Child wealth is measured in 2006. Parental wealth is measured in 1999 except in Column (4) where biological parental wealth is measured in 2006. All parents are alive in 1999 and at least one adoptive parent is alive in 2006. In Column (4), we require that both biological parents are alive in 2006. Parental wealth is calculated as combined wealth of the mother and father divided by two. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors clustered by adoptive family.

Appendix Table 1: Incidence of Zero and Negative Net Wealth

	Obs	obs with zero net wealth	obs with negative net wealth
Own-birth Children	1,200,835	11,706 (1%)	312,645 (26%)
Parents of Own-birth Children	1,200,835	9,664 (0.8%)	134,034 (11.2%)
Adoptees	2,519	38 (1.5%)	809 (32.1%)
Bio parents of adoptees	2,519	32 (1.3)%	783 (31.1%)
Adoptive parents of adoptees	2,519	24 (1%)	136 (5.4%)

Appendix Table 2: Using Various Measures of Wealth

	Inverse Hyperbolic Sine				Levels			
	Full Sample		Trim Bottom/Top 5%		Full Sample		Trim Bottom/Top 5%	
	Biological Children	Adoptees	Biological Children	Adoptees	Biological Children	Adoptees	Biological Children	Adoptees
IHS Biological Parental Wealth	0.274 (0.001)***	0.100 (0.022)***	0.286 (0.002)***	0.103 (0.029)***				
IHS Adoptive Parental Wealth		0.254 (0.042)***		0.313 (0.076)***				
Biological Parental Wealth					0.397 (0.001)***	0.068 (0.033)**	0.608 (0.003)***	0.189 (0.066)***
Adoptive Parental Wealth						0.408 (0.053)***		0.441 (0.071)***
Observations	1,080,842	2,519	1,200,833	1,971	1,200,833	2,519	1,080,842	1,971
R squared	0.056	0.085	0.086	0.084	0.086	0.358	0.053	0.130

Notes: All specifications include cohort dummies for parents and children. Parental wealth is measured in 1999 and child wealth is measured in 2006. All parents are alive in 1999 and at least one of the (adoptive) parents of (adopted) biological children is alive in 2006. Parental wealth is calculated as combined wealth of the mother and father divided by two. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors clustered by adoptive family.



Appendix Table 3: Robustness Checks  
Dependent Variable: Child Rank in Within-Cohort Wealth Distribution

VARIABLES	(1) Baseline	(2) County Dummies	(3) Nature-Nurture	(4) Wealth averaged over 3 years
Rank Biological Parental Wealth	0.132 (0.026)***	0.130 (0.026)***	0.083 (0.061)	0.128 (0.026)***
Rank Adoptive Parental Wealth	0.229 (0.027)***	0.225 (0.027)***	0.198 (0.045)***	0.243 (0.026)***
Rank Biological Parental Wealth* Rank Adoptive Parental Wealth			0.088 (0.103)	
Observations	1,971	1,971	1,971	1,963
R-squared	0.116	0.122	0.116	0.115
Child County of Residence	NO	YES	NO	NO
Parents County of Residence	NO	YES	NO	NO

Notes: The top and bottom 5 percent of parental wealth distribution have been dropped. All specifications include cohort dummies for parents and children. In Columns (1)-(3), parental wealth is measured in 1999 and child wealth is measured in 2006. In Column (4), parental wealth is averaged over 1999-2001 and child wealth is averaged over 2004-2006. All parents are alive in 1999 and at least one adoptive parent is alive in 2006. Parental wealth is calculated as combined wealth of the mother and father divided by two.

\*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors clustered by adoptive family.

Appendix Table 4: Controlling for Parental and Child Earnings and Education  
Dependent Variable: Rank of Child Net Wealth

VARIABLES	(1)	(2)	(3)	(4)
Rank Biological Parental Wealth	0.132 (0.026)***	0.133 (0.026)***	0.102 (0.026)***	0.105 (0.026)***
Rank Adoptive Parental Wealth	0.229 (0.027)***	0.231 (0.027)***	0.213 (0.026)***	0.220 (0.027)***
Observations	1,971	1,971	1,971	1,971
R-squared	0.116	0.117	0.142	0.144
Control for Adoptive Parents' Earnings	NO	YES	NO	YES
Control for Adoptive Parents' Education	NO	YES	NO	YES
Control for Children's Earnings	NO	NO	YES	YES
Control for Children's Education	NO	NO	YES	YES

Notes: The top and bottom 5 percent of the distribution have been dropped. All specifications include cohort dummies for parents and children. Child wealth is measured in 2006. All parents are alive in 1999 and at least one adoptive parent is alive in 2006. \*\*\* p<0.01, \*\* p<0.05, \* p<0.1. Standard errors clustered by adoptive family.

Education comes from the year 2000 Education Register and the education level of each adoptive parent is added separately. Parent's earnings are the log of average earnings between 1980 and 1999 and each parent's earnings is added separately. Child earnings are the log of the 3-year-average earnings between ages 34 and 36, dropping any year in which the child has zero earnings.