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## Article

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Article

# Retiree Mortality Forecasting: A Partial Age-Range or a Full Age-Range Model?

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**Abstract:** An essential input of annuity pricing is the future retiree mortality. From observed age-specific mortality data, modeling and forecasting can take place in two routes. On the one hand, we can first truncate the available data to retiree ages and then produce mortality forecasts based on a partial age-range model. On the other hand, with all available data, we can first apply a full age-range model to produce forecasts and then truncate the mortality forecasts to retiree ages. We investigate the difference in modeling the logarithmic transformation of the central mortality rates between a partial age-range and a full age-range model, using data from mainly developed countries in the Human Mortality Database (2020). By evaluating and comparing the short-term point and interval forecast accuracies, we recommend the first strategy by truncating all available data to retiree ages and then produce mortality forecasts. However, when considering the long-term forecasts, it is unclear which strategy is better since it is more difficult to find a model and parameters that are optimal. This is a disadvantage of using methods based on time-series extrapolation for long-term forecasting. Instead, an expectation approach, in which experts set a future target, could be considered, noting that this method has also had limited success in the past.

**Keywords:** age-period-cohort; Lee–Carter model with Poisson error; Lee–Carter model with Gaussian error; Plat model

## 1. Introduction

Improving human survival probability contributes greatly to an aging population. To guarantee one individual's financial income in retirement, a policyholder may purchase a fixed-term or lifetime annuity. A fixed-term or lifetime annuity is a contract offered by insurers guaranteeing regular payments in exchange for an initial premium. Since an annuity depends on survival probabilities and interest rates, pension funds and insurance companies are more likely to face a risk of longevity. Longevity risk is a potential systematic risk attached to the increasing life expectancy of policyholders, which can eventually result in a higher payout ratio than expected (Crawford et al. 2008). The concerns about longevity risk have led to a surge of interest in modeling and forecasting age-specific mortality rates.

Many models for forecasting age-specific mortality indicators have been proposed in demographic literature (see Booth and Tickle 2008, for reviews). Of these, Lee and Carter (1992) implemented a principal component method to model the logarithm of age-specific mortality rates ( $m_{x,t}$ ) and extracted a single time-varying index representing the trend in the level of mortality, from which the forecasts are obtained by a random walk with drift. Since then, the Lee–Carter (LC) method has been extended and modified (see Booth and Tickle 2008; Pitacco et al. 2009; Shang and Haberman 2018; Shang et al. 2011, for reviews). The LC method has been applied to many countries, including Belgium (Brouhns et al. 2002), Austria

(Carter and Prkawetz 2001), England and Wales (Renshaw and Haberman 2003), and Spain (Debón et al. 2008; Guillen and Vidiella-i-Anguera 2005).

Many statistical models focus on time-series extrapolation of past trends exhibited in age-specific mortality rates (see, e.g., Booth and Tickle 2008). We consider two modeling strategies: On the one hand, we can first truncate all available data to retiree ages and then produce mortality forecasts (see, e.g., Cairns et al. 2006 2009). On the other hand, we can first use the available data to produce forecasts and then truncate the mortality forecasts to retiree ages (see, e.g., Shang and Haberman 2017). In this paper, our contribution is to investigate the difference in modeling the logarithmic transformation of the central mortality rates between a partial age-range and a full age-range model, using data from mainly developed countries in the Human Mortality Database (2020). By evaluating and comparing the short-term point and interval forecast accuracies, we recommend the first strategy by truncating all available data to retiree ages and then produce mortality forecasts. However, when we consider the long-term forecasts, it is unclear which strategy is better since it is more difficult to find a model and parameters that are optimal. This is a disadvantage of using methods based on time-series extrapolation for long-term forecasting. Instead, an expectation approach, in which experts set a future target, could be considered, noting that this method has also had limited success in the past (Booth and Tickle 2008). Our recommendations could be useful to actuaries for choosing a better modeling strategy and more accurately pricing a range of annuity products.

The article is organized as follows: In Section 2, we describe the mortality data sets of 19 mainly developed countries. We revisit five time-series extrapolation models for forecasting age-specific mortality rates, which have been shown in the literature to work well across the full age range for some data sets (for more details, consult Shang 2012; Shang and Haberman 2018). Using these models as a testbed, we compare point and interval forecast accuracies between the two modeling strategies and provide our recommendations in Section 3. Conclusions are presented in Section 4.

## 2. Data Sets

The data sets used in this study were taken from the Human Mortality Database (2020). For each sex in a given calendar year, the mortality rates obtained by the ratio between “number of deaths” and “exposure to risk” are arranged in a matrix for age and calendar year. Nineteen countries, mainly developed countries, were selected, and thus 38 sub-populations of age- and sex-specific mortality rates were obtained for all analyses. The 19 countries selected all have reliable data series commencing at/before 1950. Due to possible structural breaks (i.e., two world wars), we truncate all data series from 1950 onwards. The omission of Germany is because the Human Mortality Database (2020) for a reunited Germany only dates back to 1990. The selected countries and their abbreviations are shown in Table 1, along with their last year of available data (recorded in April 2019). To avoid fluctuations at older ages, we consider ages from 0 to 99 in a single year of age and the last age group is from 100 onwards. Should we consider all ages from 0 to 110+, and we may encounter the missing-value issue and observe mortality rates outside the range of  $[0, 1]$  for some years.

**Table 1.** The 19 countries with the initial year of 1950 and their ending year listed below.

Country	Abbreviation	Last Year	Country	Abbreviation	Last Year
Australia	AUS	2014	Norway	NOR	2014
Belgium	BEL	2015	Portugal	PRT	2015
Canada	CAN	2011	Spain	SPA	2016
Denmark	DEN	2016	Sweden	SWE	2016
Finland	FIN	2015	Switzerland	SWI	2016
France	FRA	2016	Scotland	SCO	2016
Italy	ITA	2014	England & Wales	EW	2016
Japan	JPN	2016	Ireland	IRE	2014
Netherland	NET	2016	United States of America	USA	2016
New Zealand	NZ	2013			

### 3. Results

#### 3.1. Forecast Evaluation

We present 19 countries that begin in 1950 and end in the last year listed in Table 1. We keep the last 30 observations for forecasting evaluation, while the remaining observations are treated as initial fitting observations, from which we produce the one-step-ahead to 30-step-ahead forecasts. Via an expanding window approach (see also Zivot and Wang 2006), we re-estimate the parameter in the time-series forecasting models by increasing the fitted observations by one year and produce the one-step-ahead to 29-step-ahead forecasts. We iterate this process by increasing the sample size by one year until the end of the data period. The process produces 30 one-step-ahead forecasts, 29 two-step-ahead forecasts, . . . , one 30-step-ahead forecast. We compare these forecasts with the holdout samples to determine the out-of-sample forecast accuracy.

#### 3.2. Forecast Error Criteria

To evaluate the point forecast accuracy, we consider the mean absolute percentage error (MAPE) and root mean squared percentage error (RMSPE). The MAPE and RMSPE criteria measure how close the forecasts compare with the actual values of the variable being forecast, regardless of the error sign. The MAPE and RMSPE criteria can be expressed as:

$$\text{MAPE}_h = \frac{1}{p \times (31 - h)} \sum_{j=1}^{(31-h)} \sum_{x=1}^p \left| \frac{m_{x,j} - \hat{m}_{x,j}}{m_{x,j}} \right| \times 100, \quad x = 1, \dots, p,$$

$$\text{RMSPE}_h = \sqrt{\frac{1}{p \times (31 - h)} \sum_{j=1}^{(31-h)} \sum_{x=1}^p \left( \frac{m_{x,j} - \hat{m}_{x,j}}{m_{x,j}} \right)^2} \times 100$$

where  $m_{x,j}$  represents the actual holdout sample for age  $x$  in the forecasting year  $j$ ,  $p$  denotes the total number of ages, and  $\hat{m}_{x,j}$  represents the forecasts for the holdout sample.

To evaluate the pointwise interval forecast accuracy, we consider the interval score criterion of Gneiting and Raftery (2007). We consider the common case of the symmetric  $100(1 - \alpha)\%$  prediction intervals, with lower and upper bounds that were predictive quantiles at  $\alpha/2$  and  $1 - \alpha/2$ , denoted by  $\hat{m}_{x,j}^{\text{lb}}$  and  $\hat{m}_{x,j}^{\text{ub}}$ . As defined by Gneiting and Raftery (2007), a scoring rule for evaluating the pointwise interval forecast accuracy at time point  $j$  is

$$S_\alpha \left( \hat{m}_{x,j}^{\text{lb}}, \hat{m}_{x,j}^{\text{ub}}; m_{x,j} \right) = \left( \hat{m}_{x,j}^{\text{ub}} - \hat{m}_{x,j}^{\text{lb}} \right) + \frac{2}{\alpha} \left( \hat{m}_{x,j}^{\text{lb}} - m_{x,j} \right) \mathbb{1} \left\{ m_{x,j} < \hat{m}_{x,j}^{\text{lb}} \right\} \\ + \frac{2}{\alpha} \left( m_{x,j} - \hat{m}_{x,j}^{\text{ub}} \right) \mathbb{1} \left\{ m_{x,j} > \hat{m}_{x,j}^{\text{ub}} \right\},$$

where  $\mathbb{1}\{\cdot\}$  denotes the binary indicator function. The optimal interval score is achieved when  $m_{x,j}$  lies between  $\hat{m}_{x,j}^{\text{lb}}$  and  $\hat{m}_{x,j}^{\text{ub}}$ , with the distance between the upper and lower bounds being minimal. To obtain summary statistics of the interval score, we take the mean interval score across different ages and forecasting years. The mean interval score can be expressed as:

$$\bar{S}_{\alpha,h} = \frac{1}{p \times (31 - h)} \sum_{j=1}^{(31-h)} \sum_{x=1}^p S_\alpha \left( \hat{m}_{x,j}^{\text{lb}}, \hat{m}_{x,j}^{\text{ub}}; m_{x,j} \right).$$

#### 3.3. Comparison of Point and Interval Forecast Errors

Modeling and forecasting mortality can be taken place in two routes. On the one hand, we can first truncate the available data to certain ages, such as from 60 to 99 in a single year of age and 100+ as the last age, and then produce mortality forecasts for these retiree ages. On the other hand, we can first

use the available data, i.e., age-specific mortality from 0 to 99 in a single year of age and 100+ as the last age, to produce forecasts for these 101 ages and then truncate the mortality forecasts to certain ages, such as 60 to 99 in a single year of age and 100+ as the last age.

We study five time-series extrapolation models for forecasting age-specific mortality, which have been shown in the literature to work well across the full age range for some data sets. Please note that the Cairns-Blake-Dowd suite of models are not included in this paper, because they are designed just for ages 55 and over. The models that we have considered are subjective and far from extensive, but they suffice to serve as a testbed for comparing the forecast accuracy. These five models are: the Lee–Carter model with Poisson errors (see, e.g., Brouhns et al. 2002; Renshaw and Haberman 2003 2006), the Lee–Carter model with Gaussian errors (see, e.g., Booth et al. 2002; Koissi et al. 2006; Renshaw and Haberman 2003), age-period-cohort model (see, e.g., Renshaw and Haberman 2006) and the Plat model (Plat 2009).

For the short-term forecast horizon (i.e., the one-step-ahead forecast horizon), we compute the mean of the MAPEs and RMSPEs to evaluate the point forecast accuracy. From Table 2, there is an advantage of directly modeling and forecasting the truncated series for the female mortality. For modeling the male mortality, the advantage of directly modeling and forecasting the truncated series intensifies. By comparing the mean errors of the 19 countries, Table 2 shows that the most accurate forecasting method is the Plat model for providing best estimates of the female and male mortality forecasts. The Lee–Carter model with Poisson errors produces smaller MAPEs and RMSPEs than the Lee–Carter model with Gaussian errors. With the Lee–Carter model with Gaussian errors, it is advantageous to consider two components rather than only the first component.

**Table 2.** For the one-step-ahead forecast horizon  $h = 1$ , we compute the MAPE and RMSPE for each country and each model. The most accurate model for each country is highlighted in bold. For each model, we consider modeling the data with either a partial age range (termed as Partial) or a full age range (termed as Full).

Error	Country	LC (Poisson)		LC (Gaussian)		LC <sub>2</sub> (Gaussian)		APC		Plat	
		Full	Partial	Full	Partial	Full	Partial	Full	Partial	Full	Partial
MAPE	Female										
	AUS	4.45	<b>4.38</b>	6.01	4.97	4.69	4.42	7.32	6.38	4.83	4.66
	BEL	4.92	4.90	6.08	5.01	5.84	5.04	7.85	6.24	<b>4.44</b>	4.74
	CAN	3.29	3.22	3.94	3.57	3.74	3.03	4.98	4.29	3.20	<b>2.73</b>
	DEN	7.05	7.15	8.25	6.90	7.74	6.36	7.58	7.15	5.31	<b>4.79</b>
	FIN	6.43	6.13	11.56	6.80	8.26	6.80	9.31	7.94	<b>5.63</b>	5.73
	FRA	4.28	4.47	4.97	4.47	3.96	3.80	7.67	6.16	4.22	<b>3.65</b>
	ITA	3.25	3.15	8.36	3.71	<b>3.12</b>	3.17	6.49	5.56	3.68	3.43
	JPN	5.99	6.02	23.08	7.03	3.84	<b>3.23</b>	6.00	5.02	4.04	3.31
	NET	3.62	<b>3.61</b>	4.58	3.67	4.39	3.65	7.53	6.29	4.19	3.78
	NZ	<b>7.70</b>	7.75	8.65	7.90	8.33	7.85	10.02	9.38	7.75	7.74
	NOR	<b>5.25</b>	5.28	6.92	5.54	6.41	5.41	8.18	7.58	5.66	5.56
	PRT	6.06	<b>5.42</b>	11.68	5.84	7.82	5.78	8.75	7.30	6.13	6.58
	SPA	5.15	5.04	12.99	5.73	6.03	<b>4.19</b>	7.92	6.24	4.80	5.21
	SWE	<b>4.03</b>	<b>4.03</b>	4.31	4.09	4.55	4.11	8.19	7.00	4.91	4.31
	SWI	4.69	4.72	5.99	<b>4.65</b>	5.89	4.84	8.94	7.39	5.28	5.03
	SCO	6.79	6.83	8.78	6.55	7.62	6.49	7.21	6.97	5.48	<b>5.15</b>
	EW	4.79	4.72	6.27	4.49	5.20	3.51	6.32	5.80	3.92	<b>3.15</b>
	IRE	8.45	<b>7.33</b>	15.01	8.27	9.87	8.01	9.52	9.52	9.02	9.19
USA	3.33	3.22	3.84	3.20	3.81	2.37	4.50	4.28	3.43	<b>2.13</b>	
	Mean	5.24	5.12	8.49	5.39	5.85	4.85	7.59	6.66	5.05	<b>4.78</b>

Table 2. Cont.

Error	Country	LC (Poisson)		LC (Gaussian)		LC <sub>2</sub> (Gaussian)		APC		Plat	
		Full	Partial	Full	Partial	Full	Partial	Full	Partial	Full	Partial
	<u>Male</u>										
	AUS	5.39	<b>4.84</b>	7.85	5.39	5.74	5.58	7.71	5.98	5.15	5.27
	BEL	7.86	6.11	10.53	9.74	8.43	7.43	6.29	5.44	5.31	<b>4.86</b>
	CAN	5.54	4.28	7.58	5.05	5.85	4.76	5.19	4.25	4.12	<b>3.31</b>
	DEN	8.19	6.67	10.29	10.51	7.77	8.33	7.77	6.53	5.78	<b>5.52</b>
	FIN	8.12	7.49	11.84	8.57	10.02	8.43	8.11	7.10	6.61	<b>6.59</b>
	FRA	3.71	<b>3.17</b>	4.42	3.96	4.11	3.66	7.54	6.02	4.16	3.32
	ITA	5.85	4.21	11.29	5.96	4.63	3.73	5.16	4.96	3.74	<b>2.87</b>
	JPN	3.95	3.76	9.58	4.31	3.78	4.13	4.92	4.61	3.59	<b>3.27</b>
	NET	8.02	5.58	12.65	7.75	7.23	5.37	5.83	4.79	5.24	<b>4.03</b>
	NZ	9.21	8.17	12.63	10.86	9.87	10.90	9.78	8.79	<b>8.12</b>	8.38
	NOR	8.66	6.37	13.61	6.84	7.21	6.85	7.42	6.59	6.07	<b>6.02</b>
	PRT	6.12	<b>5.57</b>	8.70	7.26	8.46	6.30	7.44	6.90	5.85	5.88
	SPA	4.39	<b>3.88</b>	8.48	5.44	7.10	4.38	6.52	5.80	4.11	4.29
	SWE	5.05	4.51	9.82	5.38	6.01	5.28	8.07	6.02	4.53	<b>4.08</b>
	SWI	5.62	5.37	6.51	5.76	6.86	5.99	7.71	6.55	<b>5.27</b>	5.30
	SCO	7.06	6.20	10.63	8.56	8.42	7.51	7.51	6.71	6.29	<b>5.80</b>
	EW	4.78	4.04	8.52	5.15	5.99	4.28	6.73	4.93	3.49	<b>3.43</b>
	IRE	13.24	<b>9.34</b>	17.28	12.43	12.59	10.55	10.07	9.84	9.86	9.71
	USA	4.65	3.75	5.90	4.33	5.62	3.31	5.79	4.26	3.71	<b>2.42</b>
	Mean	6.60	5.44	9.90	7.01	7.14	6.14	7.14	6.11	5.26	<b>4.97</b>
RMSPE	<u>Female</u>										
	AUS	5.84	<b>5.77</b>	7.62	6.49	6.11	5.82	8.86	7.87	6.20	6.19
	BEL	6.40	6.43	7.55	6.39	7.37	6.47	9.46	7.64	<b>5.74</b>	6.30
	CAN	4.17	4.11	4.96	4.59	4.74	3.98	6.06	5.31	3.96	<b>3.50</b>
	DEN	9.36	9.70	10.05	9.23	9.55	8.41	9.32	8.83	6.67	<b>6.10</b>
	FIN	8.51	8.32	14.54	8.85	10.88	9.14	11.69	10.12	<b>7.46</b>	8.05
	FRA	5.57	5.87	6.22	5.51	5.03	<b>4.85</b>	9.15	7.53	5.33	4.94
	ITA	4.14	<b>4.01</b>	9.80	4.62	4.05	4.19	7.90	6.92	4.75	4.59
	JPN	7.13	7.54	25.03	8.26	4.78	<b>4.07</b>	7.32	6.30	5.00	4.51
	NET	4.62	<b>4.60</b>	5.74	4.73	5.50	4.68	8.86	7.55	5.23	4.88
	NZ	10.41	10.44	11.48	10.55	11.16	10.67	12.75	12.13	<b>10.29</b>	10.49
	NOR	<b>6.85</b>	6.92	8.63	7.20	8.13	7.13	10.15	9.43	7.23	7.35
	PRT	7.85	<b>7.03</b>	13.55	7.47	9.85	7.51	10.93	9.16	8.28	8.71
	SPA	6.53	6.51	14.70	7.18	7.85	<b>5.26</b>	9.59	7.60	6.22	6.86
	SWE	5.27	<b>5.26</b>	5.54	5.31	5.95	5.38	9.75	8.48	6.13	5.56
	SWI	5.99	6.03	7.55	<b>5.91</b>	7.48	6.12	10.63	8.98	6.63	6.54
	SCO	8.96	8.99	11.57	8.58	9.86	8.56	9.07	8.89	7.11	<b>6.77</b>
	EW	6.33	6.25	8.29	5.94	6.74	4.42	7.69	7.15	4.86	<b>3.96</b>
	IRE	11.10	<b>9.63</b>	18.75	10.82	12.88	10.58	12.22	12.43	11.76	12.26
	USA	4.19	4.13	4.79	4.13	4.73	3.27	5.91	5.65	4.29	<b>2.77</b>
	Mean	6.80	6.71	10.34	6.94	7.51	6.34	9.33	8.31	6.48	<b>6.33</b>
	<u>Male</u>										
	AUS	7.21	<b>6.78</b>	9.72	7.29	7.42	7.24	9.67	7.83	7.12	7.48
	BEL	9.98	8.47	13.14	13.50	10.48	10.02	8.86	7.83	7.39	<b>7.19</b>
	CAN	6.80	5.48	9.35	6.22	7.37	6.02	6.77	5.67	5.19	<b>4.42</b>
	DEN	11.05	9.23	13.50	14.33	10.49	11.59	10.20	8.75	7.97	<b>7.86</b>
	FIN	13.16	13.07	15.27	12.49	15.45	13.39	11.84	10.88	<b>10.56</b>	11.04
	FRA	4.72	<b>4.11</b>	5.60	5.01	5.12	4.75	9.21	7.68	5.41	4.45
	ITA	7.17	5.35	13.72	7.28	6.31	4.84	6.83	6.50	4.88	<b>4.04</b>
	JPN	4.91	4.78	10.98	5.45	4.77	5.17	6.20	5.93	4.45	<b>4.38</b>
	NET	9.68	7.05	15.47	9.82	8.96	6.80	7.61	6.36	5.60	<b>5.54</b>
	NZ	13.36	12.08	16.86	15.04	13.59	14.72	12.94	11.90	<b>11.72</b>	12.27

Table 2. Cont.

Error	Country	LC (Poisson)		LC (Gaussian)		LC <sub>2</sub> (Gaussian)		APC		Plat	
		Full	Partial	Full	Partial	Full	Partial	Full	Partial	Full	Partial
	NOR	11.32	8.98	17.02	9.45	9.82	9.73	9.89	8.98	<b>8.42</b>	8.62
	PRT	8.11	<b>7.59</b>	10.55	9.59	10.56	8.60	9.72	9.07	8.06	8.46
	SPA	5.70	<b>5.31</b>	9.97	7.01	9.17	5.73	8.23	7.71	5.47	5.64
	SWE	6.56	6.04	11.71	6.99	7.75	7.04	9.72	7.66	5.89	<b>5.60</b>
	SWI	7.56	7.30	8.50	7.77	9.24	8.19	10.02	8.60	<b>7.16</b>	7.34
	SCO	9.21	8.43	12.98	10.73	11.10	10.28	10.04	9.05	8.50	<b>7.97</b>
	EW	5.93	5.16	10.38	6.34	7.42	5.41	8.39	6.37	<b>4.50</b>	4.59
	IRE	19.05	<b>14.78</b>	23.17	17.45	18.24	15.44	15.49	14.90	15.19	15.23
	USA	5.59	4.82	6.95	5.30	6.66	4.08	7.87	6.14	4.78	<b>3.26</b>
	Mean	8.79	7.62	12.36	9.32	9.47	8.37	9.45	8.31	7.28	<b>7.13</b>

For the short-term forecast horizon (i.e., the one-step-ahead forecast horizon), we compute the  $\bar{S}_{\alpha=0.2}$  to evaluate the interval forecast accuracy. From Table 3, there is an advantage of directly modeling and forecasting the truncated series for both female and male mortality. By comparing the mean errors of the 19 countries, Table 3 shows that the most accurate forecasting method is the Plat model for providing the interval forecasts of both female and male mortality rates.

Table 3. For the one-step-ahead forecast horizon  $h = 1$ , we compute the mean interval score ( $\times 100$ ) for each country and each model.

Country	LC (Poisson)		LC (Gaussian)		LC <sub>2</sub> (Gaussian)		APC		Plat	
	Full	Partial	Full	Partial	Full	Partial	Full	Partial	Full	Partial
<u>Female</u>										
AUS	3.51	3.44	2.70	2.68	2.59	2.78	6.60	5.26	2.70	<b>2.53</b>
BEL	3.58	3.28	3.34	3.11	3.28	3.19	6.09	4.68	<b>2.61</b>	2.64
CAN	2.91	2.61	2.07	1.89	1.85	1.85	5.17	4.12	1.58	<b>1.28</b>
DEN	3.81	3.69	4.62	4.27	4.68	4.54	8.46	7.47	3.41	<b>2.60</b>
FIN	5.02	4.93	6.17	5.37	5.70	6.19	8.73	7.84	3.78	<b>3.42</b>
FRA	2.07	1.85	1.99	1.93	1.86	1.87	4.90	3.84	2.02	<b>1.75</b>
ITA	2.33	<b>1.89</b>	3.21	2.04	1.99	2.24	4.38	3.95	2.10	2.14
JPN	3.22	2.13	7.24	2.74	1.82	1.70	2.96	3.02	1.56	<b>1.40</b>
NET	3.20	3.08	2.76	2.43	2.71	2.52	7.58	5.99	2.49	<b>2.22</b>
NZ	5.87	5.75	4.86	4.68	4.90	5.63	9.18	8.19	5.01	<b>4.63</b>
NOR	4.36	4.29	3.38	3.20	3.48	3.23	6.92	6.19	3.23	<b>2.94</b>
PRT	5.34	4.33	5.54	4.21	4.58	4.20	5.57	4.69	4.24	<b>3.92</b>
SPA	4.00	3.32	3.85	2.59	2.91	<b>2.11</b>	4.24	3.39	2.62	2.69
SWE	3.00	2.93	2.83	2.76	2.94	3.18	8.19	6.44	2.46	<b>2.14</b>
SWI	3.09	3.02	4.30	4.14	4.34	4.57	7.12	5.11	<b>2.62</b>	2.64
SCO	4.79	4.62	4.01	3.80	4.12	4.45	7.67	7.51	4.12	<b>3.40</b>
EW	2.52	2.35	1.96	1.84	1.86	<b>1.79</b>	6.49	5.85	2.14	1.80
IRE	6.57	5.74	6.54	<b>4.93</b>	5.32	5.28	6.54	6.42	5.72	5.55
USA	3.18	2.92	2.03	1.78	1.91	1.27	5.89	5.51	2.67	<b>1.19</b>
Mean	3.81	3.48	3.86	3.18	3.31	3.29	6.46	5.55	3.00	<b>2.68</b>
<u>Male</u>										
AUS	5.90	5.53	4.75	4.44	<b>4.35</b>	4.57	11.27	7.63	4.60	4.57
BEL	8.79	7.71	7.73	8.24	6.87	7.81	9.87	8.85	5.99	<b>5.49</b>
CAN	5.31	4.78	3.86	3.39	3.62	3.31	8.68	6.42	3.55	<b>2.65</b>
DEN	8.23	7.04	7.48	8.33	7.08	7.97	11.80	9.96	6.37	<b>5.49</b>
FIN	10.36	9.95	9.13	8.63	9.65	9.15	13.67	12.30	<b>5.20</b>	7.39
FRA	3.04	<b>2.66</b>	2.91	2.85	2.99	3.05	10.62	8.20	3.32	2.73



Table 3. Cont.

Country	LC (Poisson)		LC (Gaussian)		LC <sub>2</sub> (Gaussian)		APC		Plat	
	Full	Partial	Full	Partial	Full	Partial	Full	Partial	Full	Partial
ITA	4.75	3.45	4.16	3.44	3.05	3.33	6.68	6.35	3.15	<b>3.04</b>
JPN	3.32	2.52	4.97	3.62	3.31	3.85	4.72	4.76	2.38	<b>2.22</b>
NET	7.46	6.26	5.85	5.12	5.25	4.35	8.30	6.69	5.20	<b>3.79</b>
NZ	10.51	9.45	9.53	9.11	8.47	9.39	14.37	12.24	8.62	<b>7.90</b>
NOR	8.36	7.57	6.95	<b>5.82</b>	5.92	6.16	10.21	9.11	6.39	5.88
PRT	7.45	6.40	7.32	6.97	7.56	7.54	9.24	9.16	5.87	<b>5.76</b>
SPA	5.35	4.32	4.54	3.80	4.40	3.45	7.62	7.21	3.48	<b>3.37</b>
SWE	5.36	5.05	5.21	4.44	4.83	4.62	12.01	9.02	4.11	<b>3.65</b>
SWI	6.23	6.00	6.76	6.63	7.01	7.00	11.77	8.96	5.15	<b>4.68</b>
SCO	8.06	7.35	7.96	7.37	8.05	8.14	11.02	10.03	6.73	<b>5.33</b>
EW	4.29	3.71	4.03	3.26	3.63	3.19	9.37	6.86	2.81	<b>2.79</b>
IRE	15.37	12.27	11.78	<b>10.10</b>	10.43	10.13	13.84	13.32	11.78	11.12
USA	5.26	4.75	3.25	2.98	3.13	1.91	11.19	8.33	4.12	<b>1.85</b>
Mean	7.02	6.15	6.22	5.71	5.77	5.73	10.33	8.71	5.20	<b>4.72</b>

For the long-term forecast horizon (i.e., the 30-step-ahead forecast horizon), we also compute the mean of the MAPEs and RMSPEs to evaluate the point forecast accuracy. From Table 4, it is unclear from the comparison of the point forecast errors if there is an advantage of modeling and forecasting the truncated series for the female mortality. In contrast, there is an advantage of modeling and forecasting the male mortality and forecasting the whole series and then truncating the mortality forecasts. By comparing the mean errors of the 19 countries, Table 4 shows that the most accurate forecasting method is the Lee–Carter model with Poisson errors for providing best estimates of the female mortality forecasts. The most accurate forecasting method is the APC model for providing the best estimates of the male mortality forecasts. The Lee–Carter model with Poisson errors produces smaller MAPEs and RMSPEs than the Lee–Carter model with Gaussian errors.

Table 4. For the 30-step-ahead forecast horizon  $h = 30$ , we compute the MAPE and RMSPE for each country and each model. The most accurate model for each country is highlighted in bold. For each model, we consider modeling the data with either a partial age range (termed as Partial) or a full age range (termed as Full).

Error	Country	LC (Poisson)		LC (Gaussian)		LC <sub>2</sub> (Gaussian)		APC		Plat	
		Full	Partial	Full	Partial	Full	Partial	Full	Partial	Full	Partial
MAPE	<u>Female</u>										
	AUS	18.78	20.05	19.09	20.46	18.62	21.58	<b>15.20</b>	26.11	20.53	26.99
	BEL	<b>15.19</b>	15.50	17.02	15.99	20.22	16.61	15.77	21.65	26.27	29.78
	CAN	13.21	12.72	12.56	9.93	<b>9.07</b>	10.68	17.96	16.29	22.76	10.32
	DEN	11.91	12.36	11.44	<b>10.71</b>	14.64	11.41	23.59	22.03	34.15	19.92
	FIN	<b>16.35</b>	18.28	16.79	18.37	21.59	20.92	23.38	27.27	26.10	29.23
	FRA	16.81	17.35	16.74	17.30	20.00	18.23	<b>13.80</b>	18.51	19.02	30.02
	ITA	18.55	21.17	26.00	24.22	24.16	22.90	<b>15.06</b>	32.54	23.41	48.43
	JPN	22.83	23.89	40.71	27.40	24.65	23.03	52.48	<b>19.82</b>	52.96	30.09
	NET	15.04	14.42	20.25	14.41	<b>11.70</b>	12.82	14.91	16.98	31.19	15.17
	NZ	34.61	35.91	38.23	40.11	38.56	36.21	<b>28.84</b>	41.59	35.94	44.18
	NOR	10.17	<b>9.52</b>	12.32	11.39	10.44	10.44	14.22	19.68	35.73	12.53
	PRT	41.89	34.61	53.01	40.15	52.41	45.16	<b>17.40</b>	46.54	47.46	54.60
	SPA	25.09	<b>24.33</b>	52.66	27.03	27.00	28.85	35.96	27.82	35.92	40.34
	SWE	14.54	13.73	15.85	12.50	13.50	12.55	16.02	16.26	24.55	<b>12.05</b>
	SWI	11.27	10.93	10.99	<b>10.83</b>	12.25	12.30	17.56	19.76	37.25	18.94
	SCO	16.91	18.29	17.55	19.72	19.75	20.56	<b>10.77</b>	25.86	31.53	20.71
	EW	17.24	19.41	55.25	19.25	17.46	20.11	<b>9.57</b>	34.12	25.03	34.45
	IRE	38.93	38.09	68.20	43.59	45.58	41.77	<b>19.47</b>	44.09	38.19	41.49
USA	8.90	7.51	11.13	5.98	7.75	<b>5.55</b>	16.65	13.24	12.56	7.83	
Mean		19.38	<b>19.37</b>	27.15	20.49	21.54	20.62	19.93	25.80	30.56	27.74



Table 4. Cont.

Error	Country	LC (Poisson)		LC (Gaussian)		LC <sub>2</sub> (Gaussian)		APC		Plat	
		Full	Partial	Full	Partial	Full	Partial	Full	Partial	Full	Partial
	<u>Male</u>										
	AUS	50.13	52.47	56.68	55.41	49.33	55.95	<b>38.21</b>	52.54	53.81	72.66
	BEL	58.34	54.88	65.69	92.47	58.83	85.97	<b>31.63</b>	46.42	87.24	46.11
	CAN	45.46	42.26	47.85	44.46	47.80	44.78	<b>33.70</b>	43.26	79.94	49.49
	DEN	63.45	63.01	63.36	67.67	66.87	67.44	66.01	57.71	73.57	<b>56.30</b>
	FIN	43.39	48.11	44.34	96.58	47.78	51.77	<b>20.85</b>	31.29	75.96	52.26
	FRA	<b>32.81</b>	33.83	35.81	34.77	37.00	35.36	39.09	35.86	60.49	39.45
	ITA	64.41	62.35	70.38	67.22	53.63	62.03	<b>32.11</b>	61.05	81.64	60.60
	JPN	<b>12.30</b>	12.68	20.98	12.79	12.91	12.82	39.53	20.94	32.12	15.61
	NET	79.68	76.14	78.67	79.29	65.19	71.08	<b>45.62</b>	51.89	79.88	47.61
	NZ	99.32	82.86	112.42	117.61	104.32	114.28	<b>63.10</b>	76.39	92.05	94.38
	NOR	87.50	73.97	88.62	71.43	65.64	67.35	<b>45.09</b>	66.85	70.00	60.19
	PRT	42.79	38.11	49.01	83.79	58.24	49.59	<b>18.01</b>	35.29	32.26	47.62
	SPA	25.43	23.13	58.45	29.85	27.15	30.84	<b>15.21</b>	25.02	26.29	28.47
	SWE	51.89	47.57	55.01	57.89	53.05	51.91	45.46	43.28	76.65	<b>42.27</b>
	SWI	36.00	37.46	41.73	42.27	39.01	45.48	<b>30.64</b>	46.42	50.82	55.93
	SCO	52.66	55.63	55.86	92.86	61.36	86.41	<b>38.51</b>	57.02	65.63	54.61
	EW	52.44	55.96	52.81	60.03	50.06	55.52	<b>35.42</b>	48.16	47.26	54.41
	IRE	110.20	91.02	112.71	111.84	111.82	104.52	<b>43.10</b>	97.26	63.07	93.34
	USA	24.07	24.99	26.07	26.81	26.12	28.13	<b>21.56</b>	23.20	70.21	31.07
	Mean	54.33	51.39	59.81	65.53	54.53	59.01	<b>36.99</b>	48.41	64.15	52.76
RMSPE	<u>Female</u>										
	AUS	21.74	23.11	22.17	23.51	21.82	24.60	<b>19.09</b>	28.34	24.41	30.76
	BEL	<b>18.25</b>	18.92	20.82	19.44	24.70	20.38	18.67	24.89	31.54	35.64
	CAN	15.37	14.95	14.53	12.18	<b>11.31</b>	13.01	22.76	21.38	28.37	14.13
	DEN	14.08	14.44	<b>13.33</b>	13.46	16.62	13.50	25.58	25.11	40.84	23.08
	FIN	<b>19.82</b>	22.37	20.14	22.67	26.18	25.80	27.37	30.78	33.02	35.66
	FRA	19.01	19.60	19.27	20.17	23.47	21.18	<b>17.97</b>	21.60	24.55	34.67
	ITA	21.18	23.45	28.25	26.36	26.44	25.17	<b>18.28</b>	36.15	28.19	57.57
	JPN	25.73	26.91	47.72	30.06	27.27	<b>25.71</b>	53.75	27.37	55.50	33.61
	NET	16.33	15.77	21.13	15.67	<b>13.93</b>	14.43	20.15	21.32	37.43	17.97
	NZ	44.37	45.84	46.55	49.71	46.68	45.70	<b>33.10</b>	48.79	42.29	50.98
	NOR	12.65	<b>12.45</b>	14.53	13.90	12.80	13.13	17.94	22.61	41.08	14.97
	PRT	48.09	40.40	61.17	46.28	60.66	52.06	<b>20.96</b>	54.09	52.80	62.86
	SPA	30.73	<b>29.51</b>	59.17	33.96	33.89	36.16	38.24	30.74	42.10	45.40
	SWE	16.23	15.50	17.54	14.61	17.18	14.73	22.05	22.21	31.04	<b>13.67</b>
	SWI	13.97	13.52	13.60	<b>13.42</b>	15.17	17.27	23.51	24.38	42.62	21.26
	SCO	22.00	24.51	20.56	25.95	22.19	26.98	<b>14.05</b>	29.29	35.82	24.81
	EW	21.84	24.08	61.19	23.89	21.62	24.44	<b>11.78</b>	38.79	29.59	39.16
	IRE	44.06	43.59	73.27	48.92	50.75	47.72	<b>24.33</b>	57.14	42.55	47.34
	USA	10.10	8.70	12.13	7.02	9.11	<b>6.77</b>	18.45	16.32	15.29	8.77
	Mean	<b>22.92</b>	23.03	30.90	24.27	25.36	24.67	23.58	30.60	35.74	32.23
	<u>Male</u>										
	AUS	58.82	61.48	66.43	64.80	57.73	66.12	<b>45.92</b>	64.35	64.19	82.67
	BEL	72.35	63.85	78.04	103.40	71.52	95.86	<b>37.45</b>	54.36	90.56	53.55
	CAN	57.57	52.42	61.40	55.54	60.60	55.81	<b>40.51</b>	52.40	84.49	57.41
	DEN	81.11	80.89	78.71	75.38	82.53	82.03	77.00	<b>66.75</b>	91.52	68.67
	FIN	51.71	57.52	52.22	111.49	55.40	61.78	<b>23.39</b>	36.30	78.51	60.90
	FRA	<b>38.15</b>	38.84	41.53	40.38	40.58	40.91	43.68	40.64	68.48	44.68
	ITA	80.70	77.32	87.79	84.15	67.36	78.15	<b>36.01</b>	71.62	85.65	75.05
	JPN	15.46	16.48	25.77	<b>15.25</b>	16.66	15.91	40.62	26.82	37.06	18.33
	NET	102.58	95.23	100.24	102.76	81.69	89.74	<b>53.64</b>	60.81	87.97	64.18
	NZ	112.83	98.95	128.69	136.49	117.87	132.75	<b>74.82</b>	91.95	102.96	107.61
	NOR	108.62	85.81	110.96	82.25	80.62	79.43	<b>50.60</b>	77.04	79.85	76.25
	PRT	49.93	44.03	56.63	95.17	64.89	55.57	<b>20.40</b>	41.52	38.67	55.07
	SPA	32.82	30.04	63.13	33.14	32.15	34.60	<b>19.17</b>	27.51	30.23	33.06
	SWE	62.19	56.65	66.37	68.98	62.41	62.39	52.86	51.01	84.63	<b>50.20</b>
	SWI	43.71	45.27	48.84	50.17	45.93	53.22	<b>36.91</b>	55.66	61.52	60.67
	SCO	64.25	66.37	68.29	105.02	72.18	97.90	<b>44.12</b>	67.61	71.81	66.18
	EW	62.45	65.98	63.72	71.74	58.80	65.56	<b>42.07</b>	58.11	56.89	63.73
	IRE	122.60	104.66	125.27	128.85	124.03	118.34	<b>49.25</b>	116.02	80.35	105.59
	USA	28.44	29.33	30.26	31.41	30.41	32.05	<b>25.85</b>	26.09	72.70	34.40
	Mean	65.59	61.64	71.28	76.65	64.39	69.37	<b>42.86</b>	57.19	72.00	62.01

For the long-term forecast horizon (i.e., the 30-step-ahead forecast horizon), we also compute the  $\bar{S}_{\alpha=0.2}$  to evaluate the interval forecast accuracy. From Table 5, there is a slight advantage of directly modeling and forecasting the truncated series for the female mortality. For modeling the male mortality, there is an advantage of modeling and forecasting the whole series and then truncating the mortality forecasts. By comparing the mean errors of the 19 countries, Table 5 shows that the most accurate forecasting method is the Lee–Carter model with Poisson errors for providing the interval forecasts of the female mortality rates. The most accurate forecasting method is the APC model for providing the interval forecasts of the male mortality rates. The Lee–Carter model with Poisson errors produces smaller MAPEs than the Lee–Carter model with Gaussian errors.

**Table 5.** For the 30-step-ahead forecast horizon  $h = 30$ , we compute the mean interval score ( $\times 100$ ) for each country and each model.

Country	LC (Poisson)		LC (Gaussian)		LC <sub>2</sub> (Gaussian)		APC		Plat	
	Full	Partial	Full	Partial	Full	Partial	Full	Partial	Full	Partial
<u>Female</u>										
AUS	8.25	<b>7.20</b>	9.30	8.94	7.75	8.80	13.42	10.46	12.31	9.29
BEL	5.99	<b>4.99</b>	6.85	6.18	7.83	10.10	8.24	11.80	19.53	12.75
CAN	12.64	10.76	7.54	8.09	5.11	7.09	23.15	21.93	23.71	<b>4.47</b>
DEN	<b>4.81</b>	5.23	5.37	5.07	6.35	18.66	<b>9.72</b>	8.85	28.91	8.85
FIN	12.80	9.57	<b>9.23</b>	9.29	12.38	162.23	25.85	20.57	27.49	16.37
FRA	5.39	<b>4.51</b>	5.49	7.33	5.45	5.69	11.37	11.82	14.18	9.99
ITA	8.81	<b>4.54</b>	12.62	10.21	6.16	8.46	6.16	7.09	7.25	13.39
JPN	6.73	<b>4.30</b>	12.47	10.43	5.40	6.06	27.24	20.42	39.63	7.77
NET	9.94	9.09	10.61	9.66	<b>6.96</b>	8.11	22.06	23.59	42.29	8.94
NZ	8.75	8.81	9.61	10.39	16.15	82.21	<b>7.78</b>	8.99	9.13	13.52
NOR	6.81	7.82	5.39	6.41	<b>5.21</b>	7.68	14.47	13.55	44.89	7.80
PRT	20.63	10.59	26.58	13.21	23.64	18.53	8.51	<b>8.49</b>	53.85	22.47
SPA	16.70	13.35	18.08	16.12	12.87	9.29	15.95	12.68	32.88	<b>8.74</b>
SWE	9.66	8.88	9.47	7.85	<b>6.16</b>	13.71	25.97	26.32	35.74	6.55
SWI	3.67	<b>3.39</b>	5.30	5.08	6.58	42.74	22.61	18.84	31.94	14.23
SCO	8.36	<b>6.08</b>	9.65	7.65	10.37	30.24	7.13	6.64	26.02	11.29
EW	<b>3.84</b>	4.04	11.42	4.45	4.81	6.20	5.09	7.44	13.97	8.55
IRE	17.25	12.39	20.57	16.64	18.89	21.43	<b>8.06</b>	9.81	15.26	20.33
USA	4.11	2.50	3.11	<b>1.80</b>	2.52	2.70	15.66	13.93	4.47	4.05
Mean	9.22	<b>7.27</b>	10.46	8.67	8.98	24.73	14.65	13.85	25.45	11.02
<u>Male</u>										
AUS	15.97	12.78	18.91	15.69	13.62	23.85	<b>10.15</b>	10.18	16.51	21.01
BEL	22.28	18.18	29.43	47.28	197.68	274.40	<b>9.25</b>	12.30	39.35	17.42
CAN	12.99	<b>11.43</b>	12.53	11.80	12.94	13.63	13.50	13.63	35.64	14.97
DEN	19.74	19.28	21.72	30.66	441.92	463.81	22.09	16.40	23.17	<b>15.61</b>
FIN	16.45	16.98	15.54	33.36	26.11	43.94	<b>8.54</b>	8.80	34.82	23.53
FRA	10.58	<b>8.20</b>	12.85	12.74	14.71	14.93	10.69	9.38	16.70	14.70
ITA	17.34	16.24	19.64	17.47	12.56	17.47	<b>7.89</b>	15.92	27.76	17.58
JPN	<b>3.58</b>	3.69	8.70	4.73	7.27	23.37	30.41	25.32	38.02	10.07
NET	23.74	21.63	24.04	20.78	18.91	23.43	14.96	<b>10.42</b>	28.78	15.71
NZ	45.00	26.13	54.39	41.50	52.59	95.29	18.23	<b>16.61</b>	28.10	42.92
NOR	26.45	25.79	25.86	27.40	16.89	23.49	14.77	<b>13.63</b>	22.20	16.68
PRT	19.87	13.68	24.01	37.26	30.92	118.32	<b>8.50</b>	12.32	12.26	22.79
SPA	27.44	14.80	35.05	19.88	15.81	14.26	<b>8.92</b>	10.80	11.09	10.42
SWE	17.71	13.46	18.53	17.60	19.87	20.01	14.96	11.34	25.19	<b>10.65</b>
SWI	9.37	9.05	14.48	13.32	28.57	28.63	<b>9.04</b>	11.19	14.00	22.49
SCO	20.85	19.01	21.61	43.03	84.51	96.68	<b>10.22</b>	12.77	25.07	18.27
EW	15.24	14.20	16.68	21.06	17.35	17.39	<b>7.48</b>	11.17	14.08	17.50
IRE	70.26	37.26	73.40	48.79	69.77	109.01	<b>12.71</b>	25.21	28.22	36.51
USA	10.95	9.65	10.60	<b>9.26</b>	10.47	10.34	12.59	11.09	42.64	11.69
Mean	21.36	16.39	24.10	24.93	57.50	75.38	<b>12.89</b>	13.60	25.45	18.97

#### 4. Conclusions

We consider forecasting retiree mortality using two modeling strategies. On the one hand, we can first truncate the available data to retiree ages and then produce mortality forecasts. On the other

hand, we can first use the available data to produce forecasts and then truncate the mortality forecasts to retiree ages. Using the empirical data from [Human Mortality Database \(2020\)](#), we investigate the short-term and long-term point and interval forecast accuracies. Between the two model strategies, we recommend the first strategy by truncating all available data to retiree ages and then produce short-term mortality forecasts. Our recommendations could be useful to actuaries for choosing a better modeling strategy and more accurately pricing a range of annuity products.

For the long-term mortality forecasts, we recommend the first strategy for modeling female mortality but the second strategy for modeling male mortality. It is difficult to recommend a strategy when the model and its parameters may not be optimal for the long-term forecasts. This is a disadvantage of using methods based on time-series extrapolation for long-term forecasting. Instead, an expectation approach, in which experts set a future target, could be considered, noting that this method has also had limited success in the past ([Booth and Tickle 2008](#)).

There are several ways in which the present study can be further extended, and we briefly mention two: (1) We could consider some machine learning methods to model mortality forecasts (see, e.g., [Perla et al. 2020](#); [Richman and Wüthrich 2020](#)). (2) The results depend on the age range considered as well as the selected countries. The R code for reproducing the results can be provided upon request from the corresponding author.

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## References

- Booth, Heather, and Leonie Tickle. 2008. Mortality modelling and forecasting: A review of methods. *Annals of Actuarial Science* 3: 3–43. [[CrossRef](#)]
- Booth, Heather, John Maindonald, and Len Smith. 2002. Applying Lee-Carter under conditions of variable mortality decline. *Population Studies* 56: 325–36. [[CrossRef](#)] [[PubMed](#)]
- Brouhns, Natacha, Michel Denuit, and Jeroen K. Vermunt. 2002. A Poisson log-bilinear regression approach to the construction of projected life-tables. *Insurance: Mathematics and Economics* 31: 373–93.
- Cairns, Andrew J. G., David Blake, and Kevin Dowd. 2006. A two-factor model for stochastic mortality with parameter uncertainty: Theory and calibration. *Journal of Risk and Insurance* 73: 687–718. [[CrossRef](#)]
- Cairns, Andrew J. G., David Blake, Kevin Dowd, Guy D. Coughlan, David Epstein, Alen Ong, and Igor Balevich. 2009. A quantitative comparison of stochastic mortality models using data from England and Wales and the United States. *North American Actuarial Journal* 13: 1–35. [[CrossRef](#)]
- Carter, Lawrence R., and Alexia Prskawetz. 2001. *Examining Structural Shifts in Mortality Using the Lee-Carter Model*; Technical Report. Rostock: Center for Demography and Ecology Information, University of Wisconsin-Madison.
- Crawford, Thomas, Richard de Haan, and Chad Runchey. 2008. *Longevity Risk Quantification and Management: A Review of rElevant Literature*; Technical Report. Schaumburg: Society of Actuaries.
- Debón, Ana, F. Montes, and Francisco Puig. 2008. Modelling and forecasting mortality in Spain. *European Journal of Operational Research* 198: 624–37. [[CrossRef](#)]
- Gneiting, Tilmann, and Adrian E. Raftery. 2007. Strictly proper scoring rules, prediction and estimation. *Journal of the American statistical Association* 102: 359–78. [[CrossRef](#)]
- Guillen, Montserrat, and Antoni Vidiella-i-Anguera. 2005. Forecasting Spanish natural life expectancy. *Risk Analysis: An International Journal* 25: 61–1170. [[CrossRef](#)] [[PubMed](#)]

- Human Mortality Database. 2020. University of California, Berkeley (USA), and Max Planck Institute for Demographic Research (Germany). Available online: [www.mortality.org](http://www.mortality.org) or [www.humanmortality.de](http://www.humanmortality.de) (accessed on 19 December 2019).
- Koissi, Marie-Claire, Arnold F. Shapiro, and Göran Högnäs. 2006. Evaluating and extending the Lee-Carter model for mortality forecasting: Bootstrap confidence interval. *Insurance: Mathematics and Economics* 38: 1–20. [[CrossRef](#)]
- Lee, Ronald D., and Lawrence R. Carter. 1992. Modeling and forecasting U.S. mortality. *Journal of the American Statistical Association* 87: 659–71. [[CrossRef](#)]
- Perla, Francesca, Ronald Richman, Salvatore Scognamiglio, and Mario V. Wuthrich. 2020. *Time-Series Forecasting of Mortality Rates Using Deep Learning*; Technical Report. Amsterdam: SSRN.
- Pitacco, Ermanno, Michel Denuit, Steven Haberman, and Annamaria Olivieri. 2009. *Modelling Longevity Dynamics for Pensions and Annuity Business*. Oxford: Oxford University Press.
- Plat, Richard. 2009. On stochastic mortality modeling. *Insurance: Mathematics and Economics* 45: 393–404.
- Renshaw, Arthur E., and Steven Haberman. 2003. Lee-Carter mortality forecasting with age-specific enhancement. *Insurance: Mathematics and Economics* 33: 255–72. [[CrossRef](#)]
- Renshaw, Arthur E., and Steven Haberman. 2006. A cohort-based extension to the Lee-Carter model for mortality reduction factors. *Insurance: Mathematics and Economics* 38: 556–70. [[CrossRef](#)]
- Richman, Ronald, and Mario V. Wüthrich. 2020. A neural network extension of the Lee-Carter model to multiple populations. *Annals of Actuarial Science* in press.
- Shang, Han Lin. 2012. Point and interval forecasts of age-specific life expectancies: A model averaging approach. *Demographic Research* 27: 593–644. [[CrossRef](#)]
- Shang, Han Lin, and Steven Haberman. 2017. Grouped multivariate and functional time series forecasting: An application to annuity pricing. *Insurance: Mathematics and Economics* 75: 166–79.
- Shang, Han Lin, and Steven Haberman. 2018. Model confidence sets and forecast combination: An application to age-specific mortality. *Genus* 74: 19. [[CrossRef](#)] [[PubMed](#)]
- Shang, Han Lin, Heather Booth, and Rob J. Hyndman. 2011. Point and interval forecasts of mortality rates and life expectancy: A comparison of ten principal component methods. *Demographic Research* 25: 173–214. [[CrossRef](#)]
- Zivot, Eric, and Jiahui Wang. 2006. *Modeling Financial Time Series with S-PLUS*. New York: Springer.



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