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Working Paper

Investment in stadia and regional economic development: evidence from FIFA World Cup 2006

Hamburg Contemporary Economic Discussions, No. 16

Provided in Cooperation with:

University of Hamburg, Chair for Economic Policy

Suggested Citation: Feddersen, Arne; Grötzinger, André L.; Maennig, Wolfgang (2008) : Investment in stadia and regional economic development: evidence from FIFA World Cup 2006, Hamburg Contemporary Economic Discussions, No. 16, ISBN 978-3-940369-47-5, University of Hamburg, Faculty of Business, Economics and Social Sciences, Chair for Economic Policy, Hamburg

This Version is available at:

<https://hdl.handle.net/10419/25364>

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INVESTMENT IN STADIA AND REGIONAL ECONOMIC DEVELOPMENT— EVIDENCE FROM FIFA WORLD CUP 2006 STADIA

HAMBURG CONTEMPORARY

Urban
Transport
Media
Sports
Socio-
Regional
Real Estate
Architectural

ECONOMIC DISCUSSIONS

No. 16

Hamburg Contemporary Economic Discussions

University of Hamburg

Faculty Economics and Social Science

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ISSN 1865 - 2441 (Print)

ISSN 1865 - 7133 (Online)

ISBN 978 - 3 - 940369 - 46 - 8 (Print)

ISBN 978 - 3 - 940369 - 47 - 5 (Online)

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Investment in Stadia and Regional Economic Development – Evidence from FIFA World Cup 2006

Abstract: Using the case of the new stadiums for the FIFA World Cup 2006 in Germany, this paper is the first multivariate work that examines the potential income and employment effects of new stadiums outside of the USA. This study is also the first work on this topic that conducts tests on the basis of a (serial correlation consistent) Difference-in-Difference model with level and trends. As a robustness check, we use the “ignoring time series information” model in a form that is modified for non-synchronous interventions. We were not able to identify income or employment effects of the construction of new stadiums for the FIFA World Cup 2006, which are significantly different from zero.

Keywords: Sports Economics, Regional Economics, Stadia Infrastructure, Difference-in-Difference Model

JEL classification: H54, L83, R12, R53

Version: October 2008

1 Introduction

A series of studies on Metropolitan Statistical Areas (MSAs) in the USA revealed that new sport stadiums do not generate significant income and/or employment effects in their host cities,¹ challenging the “boosters” view of many politicians and sport officials who claim beneficial effects for the local economy (and hence, a justification for public financial support).

Using the case of the new stadiums for the FIFA World Cup 2006 in Germany, this paper is the first multivariate work that examines the potential income and employment effects of new stadiums outside of the USA. Such a study is generally interesting set against the background of the different urban structures in the USA and Europe. In addition, a non-US study is especially interesting because of decade-long US tradition of allocating the stadiums in suburban areas, whereas

¹ See BAADE (1987, 1994, 1996), BAADE & DYE (1990), BAADE & SANDERSON (1997), COATES & HUMPHREYS (1999, 2000, 2001, 2003).

European stadiums are mostly located near to the city center (FEDDERSEN & MAENNIG, 2008). NELSON (2001) argued that (US-)studies concluding insignificant effects on the home cities of stadiums are misleading, since the data are based on stadia built in the 1960s-1980s. On closer examination of the economic impact, it is evident that stadiums built in Central Business Districts (CBD) or downtown sites have a positive effect, while for suburban stadiums the effects on regional economic development are insignificant or even negative.²

This study is also the first work on this topic that conducts tests on the basis of a Difference-in-Difference (DD) model with levels and trends. To address the problem of potential serial correlation in DD models (BERTRAND, DUFLO, & MULLAINATHAN, 2004), we use a serial correlation consistent arbitrary variance-covariance matrix. As robustness check we use the “ignoring time series information” (ITSI) model in a form that is modified for non-synchronous interventions. The paper is organized as followed. Section 2 elaborates on the data, section 3 presents methods and results, and section 4 concludes.

2 Data

The FIFA World Cup 2006 in Germany was held in 12 different stadiums (Berlin, Cologne, Dortmund, Frankfurt, Gelsenkirchen, Hamburg, Hannover, Kaiserslautern, Leipzig, Munich, Nuremberg, Stuttgart). The investment costs for new construction or major renovations totaled an amount of nearly €1.6 billion for twelve stadiums (FIFA, 2006).³ Additional €1.6 billion was invested into stadia related infrastructure (BÜTTNER, MAENNIG, & MENßEN, 2005). As the aim of this analysis is to identify the effects of the FIFA World Cup stadiums, these twelve cities will be used as the treatment group in the DD model. During the period of observation, several additional stadium construction projects were undertaken in Ger-

² MELANIPHY (1996) and SANTEE (1996) also argued that stadiums in inner cities might be more efficient for the regional development of these cities.

³ Every World Cup stadium was at least renovated. The average expenditure per city was €116.7 million with a minimum investment of €36.0 million (Dortmund) and a maximum investment of €280.0 million (Munich).

many. To avoid biased results, in addition to the FIFA World Cup stadiums, all relevant stadium construction projects (including the FIFA World Cup stadiums) were used as the treatment group in a second DD regression.

Tab. 1 Relevant Stadium Construction Projects in Germany, 1996 to 2005

| City | Stadium | Capacity | Team(s) | Costs | Construction | |
|-----------------------|---------------------------------|---------------|-----------------------------|--------------|-----------------|-------------------|
| | | | | | Start | End |
| Berlin | Olympiastadion | 74,000 | Hertha BSC Berlin | 242.0 | Aug 2000 | Aug 2004 |
| Bremen | Weserstadion | 42,100 | Werder Bremen | 18.0 | May 2003 | Jul 2004 |
| Cologne | RheinEnergy-Stadion | 50,374 | 1. FC Köln | 117.5 | Jan 2002 | Jul 2004 |
| Cottbus | Stadion der Freundschaft | 22,746 | FC Energie Cottbus | 12.0 | Apr 2002 | Jul 2003 |
| Dortmund | Signal Iduna Park | 83,000 | Borussia Dortmund | 36.0 | May 2002 | Jul 2003 |
| Düsseldorf | LTU arena | 52,000 | Fortuna Düsseldorf | 218.0 | Sep 2002 | Jan 2005 |
| Duisburg | MSV-Arena | 31,514 | MSV Duisburg | 43.0 | Oct 2003 | Jan 2005 |
| Frankfurt | Commerzbank-Arena | 51,500 | Eintracht Frankfurt | 126.0 | Jul 2002 | May 2005 |
| Gelsenkirchen | Veltins-Arena | 61,524 | FC Schalke 04 | 192.0 | Nov 1998 | Jul 2001 |
| Hamburg | HSH-Nordbank-Arena | 57,000 | Hamburger SV | 97.0 | Jun 1998 | Aug 2000 |
| Hannover | AWD-Arena | 49,000 | Hannover 96 | 63.0 | Feb 2003 | Jan 2005 |
| Kaiserslautern | Fritz-Walter-Stadion | 48,500 | 1. FC Kaiserslautern | 48.3 | Aug 2004 | Apr 2006 |
| Leipzig | Zentralstadion | 44,193 | Sachsen Leipzig | 90.6 | Dec 2000 | March 2004 |
| Magdeburg | Stadion Magdeburg | 27,000 | 1. FC Magdeburg | 30.9 | March 2005 | Dec 2006 |
| Mönchengladbach | Borussia-Park | 54,057 | Borussia M'gladbach | 87.0 | Jan 2002 | Jul 2004 |
| Munich | Allianz Arena | 69,901 | FC Bayern München | 280.0 | Feb 2002 | May 2005 |
| Nuremberg | easyCredit-Stadion | 46,780 | 1. FC Nürnberg | 56.0 | Nov 2003 | Jul 2005 |
| Rostock | DKB-Arena | 30000 | FC Hansa Rostock | 55.0 | May 2000 | Aug 2001 |
| Stuttgart | Gottlieb-Daimler-Stadion | 55,896 | VfB Stuttgart | 51.6 | Jan 2004 | Jan 2006 |
| Wolfsburg | Volkswagen Arena | 29,161 | VfL Wolfsburg | 51.0 | May 2001 | Nov 2002 |

Source: SKRENTNY (2001); FIFA (2006); STADIONWELT (2007); FIFA World Cup 2006 stadia are marked in bold letters.

The analytical framework for this study comprises data of the 118 most populated large urban districts (*“Kreisfreie Städte”*) in Germany in 1995, as reported by the ARBEITSKREIS VOLKSWIRTSCHAFTLICHE GESAMTRECHNUNG DER LÄNDER (2007b).⁴ As variables for the regional economic development, the income of private households per capita (ARBEITSKREIS VOLKSWIRTSCHAFTLICHE GESAM-

⁴ See Table A1 in the annex for a complete list of the large urban districts.

TRECHNUNGEN DER LÄNDER, 2007b) as well as the number of people employed (ARBEITSKREIS VOLKSWIRTSCHAFTLICHE GESAMTRECHNUNGEN DER LÄNDER, 2007a) in these 118 large urban districts are considered.

Fig. 1. Large Urban Districts in Germany and Stadia Construction Projects



Notes: World Cup venues are marked in black, German large urban districts are marked in grey. Own illustration.

For the income of private households, the period of observation is 1995 to 2005, i.e. a time Span of 11 years; the period from 1996 to 2005, i.e. a time Span of ten years, is considered for employment data. As the data availability starts in 1995, no structural breaks due to German reunification have to be considered.⁵

Several additional indicators of the regional economic development could be considered. HOTCHKISS, MOORE, & ZOBAY (2003), for instance, suggested that the DD equation could be estimated for population. As one easily can see, a sport venue or sport franchise (sport club) might increase the attractiveness of a city from

⁵ Start and end of the observation periods are determined by data availability from EUROSTAT and VGRDL.

a resident’s point of view. As a consequence, migration into the city may occur. Thus, initially, it might be appropriate to test for a population effect. However, since it is difficult to assume that unemployed persons will migrate due to the increased attractiveness of a city, we can assume that most migrants will be working in their new city. Thus a strong correlation between population and employment exists and an additional DD analysis on population is unnecessary.

3 Method and Results

3.1 DD Model with Level and Trend

The aim of this paper is to examine if stadium construction projects in Germany – especially those of the FIFA World Cup 2006 – have a significant impact on the economic development of the regions in which they are located. For this purpose, we use a DD estimation. This is a common approach for identifying the effect of a specific intervention or treatment. Therefore, one has to compare the differences in outcome before and after an intervention for groups affected by the intervention to the difference for unaffected groups (BERTRAND, DUFLO, & MULLAINATHAN, 2004, p. 249).

We focus our interest on differences in levels and trends for two variables: employment and income. Since the stadium construction work did not start at the same point in time for all cities (see Table 1) the pre-period and the post-period are not the same for all cities of the treatment group, and they are not even defined for the control cities. Thus, in contrast to many DD models,⁶ no dummy variable for the post-period of all cities will be included. Equation (1) and (2) contain the modified DD model:

$$Z_{it} = \alpha + \beta_1 PT_{it} + \beta_2 trend + \beta_3 TT_{it} + \beta_4 PTT_{it} + \varepsilon_{it} \quad (1)$$

$$\text{with } \varepsilon_{it} = \mu_i + \nu_{it}$$

⁶ See e.g. HAGN & MAENNIG (2008a, 2008b), JASMAND & MAENNIG (2008) or HOTCHKISS, MOORE, & ZOBAY (2003) for the use of a general post period dummy.

where Z_{it} is the income of private households in city i in year t or the employment in city i in year t , respectively. α denotes the intercept term. $trend$ is a trend variable for all 118 large urban districts starting with the value of one in year 1995 (1996) and ends with a value of eleven (ten) in year 2005. No dummy variable for the treatment group is included because our model is a fixed effects model with separate dummies for all large urban districts capturing the treatment group effects. PT_{it} is a dummy for the post intervention phase of the treatment group. It takes the value of one for cities with relevant stadia construction projects from the year of the start of the construction work⁷ and zero otherwise. PTT_{it} denotes a variable that covers a post period trend for the treatment cities. It is the product of the variables PT_{it} and $trend$. In the years before the start of the construction project it takes the value of zero and afterwards it displays the corresponding value of the $trend$ variable. β_1 , β_2 , β_3 , and β_4 are coefficients to be estimated. μ_i covers the unobserved individual specific effects (fixed effects) while v_{it} denotes the remainder disturbance.

The coefficients of interests are β_1 and β_4 since they are measuring the level and trend effect of the intervention (stadium construction project) of the treatment cities. If a stadium construction project produces an impact on employment and income, then these coefficients need to be significant. Due to need for workers to accomplish the construction, the demand for employees will increase. Thus, a positive sign of the level effect (PT_{it}) could be found in the employment model.

In contrast, the trend effect on income per capita is theoretically ambiguous. If we assume that the attractiveness of a city increase in the eyes of residents and non-residents (for example, because of an eye-catching new stadium and its asso-

⁷ As the employment and income data are on a yearly basis and as the construction work does not always starts at the beginning of year, no effect could be found for a year in which a construction project starts at the year's end. To deal with this problem, stadium constructions will be considered only for a specific year if the start of work lies in first three quarters of this year. If the construction work started in the last quarter of a year, the following year will be treated as starting point.

ciated feel-good effects⁸), then migration into the city may occur. If the population increases, the labor supply might increase, potentially leading to decreasing wages (“compensating differentials”, CARLINO & COULSON, 2004).

To isolate the effect of the pure construction phase, a second variant of model (1) will be estimated:

$$Z_{it} = \alpha + \beta_1 PT_{it} + \beta_2 C + \beta_3 trend + \beta_4 TT_{it} + \beta_5 PTTC_{it} + \varepsilon_{it} \quad (2)$$

$$\text{with } \varepsilon_{it} = \mu_i + \nu_{it}$$

The variables *trend*, TT_{it} and PT_{it} are identical to those in model (1). To isolate the effects of the construction phase, the dummy variable *C* takes the value of one during the construction work and the value of zero otherwise.⁹ $PTTC_{it}$ is a post intervention trend for the treatment group that starts after the construction work has finished since we expect that changes in the growth trend will occur not due to the construction but, rather, due to advancements in the attractiveness of the city that are derived only from the completed stadium. It has to be admitted, though, that, due to data limitations, for some stadia projects (e.g. Kaiserslautern or Stuttgart) only a few observations are available for $PTTC$, making it statistically demanding to isolate any post-construction effects for these cities.

As shown by BERTRAND, DUFLO, & MULLAINATHAN (2004), DD models are frequently subject to serial correlation, which might lead to an overestimation of the significance of the “intervention” dummy. To check for such problems, we performed an LM test for serial correlation in a fixed effects model as suggested by BALTAGI (2001, pp. 94-95).¹⁰ This test is performed on the residuals of standard

⁸ See MAENNIG (2006) for an overview of the effects of iconic architecture of sporting venues and MAENNIG & PORSCHE (2008) for a first contribution dealing with the feel-good effects of large sporting events.

⁹ The periods of construction can be found in columns 6 and 7 of Table 1. As construction work is not always started at the beginning of a year, the dummy takes the value of one if the works start before October or does not end before April of the respective year.

¹⁰ The LM test statistic is $LM_5 = \sqrt{NT^2/(T-1)(\tilde{v}'\tilde{v}_{-1}/\tilde{v}'\tilde{v})}$, which is asymptotically distributed as $N(0,1)$.

fixed effects regressions of the above described models (1) and (2) for income and employment.¹¹

Tab. 2 Test for Serial Correlation

| Endogenous variable | Model (1) | | Model (2) | |
|---------------------|--------------|---------------|--------------|---------------|
| | Treatment WC | Treatment ALL | Treatment WC | Treatment ALL |
| Income | 23.411 | 23.413 | 23.249 | 23.251 |
| Employment | 33.964 | 23.186 | 23.367 | 23.371 |

Notes: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$.

The LM statistic indeed rejects the null hypothesis of no serial correlation in each case.

For such a case, BERTRAND, DUFLO, & MULLAINATHAN (2004) suggest using an arbitrary variance-covariance matrix, which is consistent in the presence of any correlation pattern within cross section over time. Table 3 and Table 4 show the regression results of the DD coefficients and the corresponding t -statistics computed using an arbitrary variance-covariance matrix.

¹¹ The “intervention” coefficients of these regressions are often significant. But in the line with BERTRAND, DUFLO, & MULLAINATHAN (2004) the estimates might be inefficient.

Tab. 3 DD Model with Fixed Effects for Income of Private Households

| | Model (1) | | Model (2) | |
|---------------------|-------------------------------------|-----------------------------------|-------------------------------------|-------------------------------------|
| | Treatment WC 12 | Treatment ALL 20 | Treatment WC 12 | Treatment ALL 20 |
| Constant | 9.551 ^{***} (3,539.662) | 9.550 (3,663.220) | 9.551 ^{***} (3,502.198) | 9.551 ^{***} (3,594.649) |
| P | -6.109e ⁻⁴ (-0.432) | 6.297e ⁻⁴ (0.733) | -0.017 (-1.257) | -0.020 [*] (-1.787) |
| C | – | – | -0.005 (-0.694) | -0.008 [*] (-1.742) |
| trend | 0.019 ^{***} (38.905) | 0.019 ^{***} (37.428) | 0.019 ^{***} (38.897) | 0.019 ^{***} (37.413) |
| TT | -0.005 (-0.465) | -0.008 (-0.820) | -8.063e ⁻⁴ (-0.528) | 2.243e ⁻⁴ (0.213) |
| PTT | 8.260e ⁻⁴ (0.201) | -7.980e ⁻⁵ (-0.018) | – | – |
| PTTC | – | – | 0.010 (1.365) | 0.009 (1.228) |
| R ² | 0.881 | 0.881 | 0.882 | 0.882 |
| adj. R ² | 0.881 | 0.881 | 0.881 | 0.881 |
| F-Stat | 441.250 ^{***} | 550.130 ^{***} | 352.928 ^{***} | 374.450 ^{***} |
| N | 118 | 118 | 118 | 118 |
| T | 11 | 11 | 11 | 11 |
| N*T | 1298 | 1298 | 1298 | 1298 |

Notes: ^{***} $p < 0.01$, ^{**} $p < 0.05$, ^{*} $p < 0.10$. *t*-statistics are in parentheses. Standard errors are computed using an arbitrary variance-covariance matrix as suggested by BERTRAND, DUFLO, & MULLAINATHAN (2004, pp. 270-272).

Tab. 4 DD Model with Fixed Effects for Employment

| | Model (1) | | Model (2) | |
|--------------------|---------------------------------|---------------------------------|---------------------------------|---------------------------------|
| | Treatment WC 12 | Treatment ALL 20 | Treatment WC 12 | Treatment ALL 20 |
| Constant | 11.244 *** (1,804.627) | 11.245 *** (1,954.357) | 11.245 *** (1,950.524) | 11.244 *** (1,868.912) |
| P | 0.006 (0.723) | 0.015 (1.106) | 0.015 (0.623) | 0.002 (0.100) |
| C | – | – | -0.009 (-1.159) | -0.010 (-1.371) |
| trend | 2.189e ⁻⁴ (0.222) | 3.647e ⁻⁴ (0.375) | 9.702e ⁻⁴ (0.003) | 2.189e ⁻⁴ (0.222) |
| TT | 0.006 (1.494) | 0.005 (1.224) | 0.004 (1.089) | 0.005 (1.413) |
| PTT | -0.008 (-1.523) | -0.006 (-1.146) | – | – |
| PTTC | – | – | -0.001 (-0.071) | -0.007 (-0.525) |
| R ² | 0.979 | 0.979 | 0.979 | 0.979 |
| Adj.R ² | 0.979 | 0.979 | 0.978 | 0.978 |
| F-Stat | 10,256.240 *** | 10,330.930 *** | 9,620.552 *** | 9,596.174 *** |
| N | 118 | 118 | 118 | 118 |
| T | 10 | 10 | 10 | 10 |
| N*T | 1180 | 1180 | 1180 | 1180 |

Notes: ***p<0.01, **p<0.05, *p<0.10. *t*-statistics are in parentheses. Standard errors are computed using an arbitrary variance-covariance matrix as suggested by BERTRAND, DUFLO, & MULLAINATHAN (2004, pp. 270-272).

In all four estimated income models, the trend variable *trend* is significant at the 1%-level, while it is not significant for the employment estimations. Not surprisingly, this means that there is a positive trend in income for all regarded 118 German large urban districts within the observation period. The treatment trend dummy is insignificant in all models, implying that there is no systematic difference between the treatment and control groups in the growth pattern of urban districts. The coefficients of the post-period dummy *PT* of the treatment urban districts and the respective coefficient of the post-trend dummy *PTT* – the objects

of interest – are insignificant for all estimations. The results are not affected by accounting for a special construction effect, as shown in model (2) of the income and employment regressions. Thus, the hypothesis of no income and employment effect of the stadia construction projects in the 12 respectively 20 urban districts with completed stadia construction cannot be rejected.

3.2 Ignoring Time Series Information DD Model

To check robustness, we will use the “ignoring time series information” (ITSI) model in its modification for non synchronous interventions (BERTRAND, DUFLO, & MULLAINATHAN, 2004, pp. 267-269). In a first step, Z_{it} (equation 1 and 2) was regressed on city fixed effects, time fixed effects and relevant covariates.¹² In the second step, the residuals of only the treatment group will be taken into account. These residuals will be divided into two groups: (1) residuals from years before the start of a stadia construction project, and (2) residuals from years after the start of a stadia construction project. The stadia effect can then be analyzed by an OLS regression of a two-period regression of the residuals from the treatment cities only. Consistent t -statistics can be obtained from this OLS regression.¹³

Tab. 5 ITSI DD Model for Income of Private Households

| | Model (1) | | Model (2) | |
|---------------------|--------------------|--------------------|--------------------|--------------------|
| | Treatment WC | Treatment ALL | Treatment WC | Treatment ALL |
| Constant | 0.046 (1.081) | 0.020 (0.596) | 0.047 (1.111) | 0.022 (0.635) |
| POST | -0.006 (-0.102) | -0.001 (-0.030) | -0.005 (-0.084) | -0.001 (-0.020) |
| R ² | 0.045 | 0.019 | 0.030 | 0.024 |
| adj. R ² | 0.001 | 0.007 | 0.014 | 0.002 |

Notes: *** $p < 0.01$, ** $p < 0.05$, * $p < 0.10$. t -statistics are in parentheses. Coefficients are from a two-step process using OLS.

¹² As done in the previous section, two different variants have been analyzed: (1) no covariates are considered; (2) only a construction dummy is considered.

¹³ As the numbers of cities is not small, the t -statistics don't have to be adjusted (BERTRAND, DUFLO, & MULLAINATHAN, 2004).

Tab. 6 ITSI DD Model for Employment

| | Model (1) | | Model (2) | |
|---------------------|----------------------|----------------------|----------------------|----------------------|
| | Treatment WC | Treatment ALL | Treatment WC | Treatment ALL |
| Constant | 1.573 *** (6.016) | 1.192 *** (5.734) | 1.577 *** (6.050) | 1.196 *** (5.760) |
| POST | -0.008 (-0.022) | -0.004 (-0.013) | -0.014 (-0.039) | -0.003 (-0.010) |
| R ² | 0.095 | 0.057 | 0.035 | 0.016 |
| adj. R ² | 0.054 | 0.032 | 0.009 | 0.010 |

Notes: ***p<0.01, **p<0.05, *p<0.10. *t*-statistics are in parentheses. Coefficients are from a two-step process using OLS.

The results of the ITSI models as shown in Table 5 and Table 6 confirm the findings of the DD model estimated in section 3.1. No coefficient in the ITSI models is significant on any conventional level. The results of the robustness check support the results from the DD model using an arbitrary variance-covariance matrix.

4 Conclusion

We were not able to identify income or employment effects of the construction of the new stadiums for the World Cup 2006, which are significantly different from zero, in the urban districts with completed new stadiums in the period leading up to and after the FIFA World Cup 2006.

We nevertheless hesitate to share the concern expressed both implicitly and explicitly in many of the comparable sports economic studies that the positive effects of new stadiums claimed by many sports protagonists are not true for three reasons. Firstly, other effects such as the feel-good benefit for the population and/or image effects that are difficult to quantify, may be sufficiently important to justify major new stadiums and/or subsidies for them via public funds. With

image effects and feel-good effects, economic empiricism in regards to sports is still in its infancy.¹⁴

Secondly, the treatment group in the selected form of municipality areas might be still too large and too highly aggregated to statistically prove significant effects. Studies on the effects of major sports venues on property values in surrounding areas indicate a maximum affect area of around 3,000 metres (AHLFELDT & MAENNIG, 2007a, 2007b; TU, 2005).

¹⁴ For the measurement of the benefit of the Olympic Games in London 2012 cf. ATKINSON *et al.* (2008); for the measurement of the willingness to pay for the Soccer World Cup 2006 (before and after the event) cf. HEYNE, MAENNIG, & SÜßMUTH (2007).

Appendix

Tab. A1. Population of the 118 largest urban districts (“kreisfreie Städte”) in Germany in 1995

| No. | City | Population in 1995 |
|-----|----------------------------|--------------------|
| 1 | Berlin | 3,471,003 |
| 2 | Hamburg | 1,707,251 |
| 3 | München | 1,240,465 |
| 4 | Köln | 964,597 |
| 5 | Frankfurt am Main | 651,097 |
| 6 | Essen | 616,340 |
| 7 | Dortmund | 599,966 |
| 8 | Stuttgart | 586,954 |
| 9 | Düsseldorf | 572,171 |
| 10 | Bremen | 549,157 |
| 11 | Duisburg | 535,473 |
| 12 | Leipzig | 524,870 |
| 13 | Hannover | 523,574 |
| 14 | Dresden | 496,863 |
| 15 | Nürnberg | 493,940 |
| 16 | Bochum | 400,608 |
| 17 | Wuppertal | 382,600 |
| 18 | Saarbrücken Stadtverband | 358,365 |
| 19 | Bielefeld | 324,115 |
| 20 | Mannheim Universitätsstadt | 313,880 |
| 21 | Gelsenkirchen | 292,061 |
| 22 | Bonn | 291,863 |
| 23 | Chemnitz | 291,331 |
| 24 | Halle (Saale) | 287,052 |
| 25 | Karlsruhe | 276,544 |
| 26 | Wiesbaden | 266,532 |
| 27 | Mönchengladbach | 266,095 |
| 28 | Münster | 264,696 |
| 29 | Magdeburg | 262,557 |
| 30 | Augsburg | 260,952 |
| 31 | Braunschweig | 253,513 |
| 32 | Krefeld | 249,821 |
| 33 | Aachen | 247,460 |
| 34 | Kiel | 246,595 |
| 35 | Rostock | 230,768 |
| 36 | Oberhausen | 224,896 |
| 37 | Lübeck Hansestadt | 216,933 |
| 38 | Hagen | 212,909 |
| 39 | Erfurt | 212,532 |
| 40 | Kassel | 201,628 |
| 41 | Freiburg im Breisgau | 198,394 |
| 42 | Mainz | 184,329 |
| 43 | Hamm | 183,734 |
| 44 | Herne | 179,973 |
| 45 | Mülheim an der Ruhr | 176,602 |
| 46 | Osnabrück | 168,106 |
| 47 | Ludwigshafen am Rhein | 167,872 |
| 48 | Solingen | 165,794 |
| 49 | Leverkusen | 162,051 |
| 50 | Oldenburg (Oldenburg) | 150,540 |
| 51 | Potsdam | 144,941 |
| 52 | Darmstadt | 138,973 |
| 53 | Heidelberg | 138,612 |
| 54 | Bremerhaven | 130,720 |
| 55 | Cottbus | 127,791 |
| 56 | Würzburg | 127,627 |
| 57 | Wolfsburg | 126,782 |
| 58 | Regensburg | 125,809 |
| 59 | Gera | 124,971 |
| 60 | Reimscheid | 122,710 |

| | | |
|-----|----------------------------|---------|
| 61 | Heilbronn | 121,745 |
| 62 | Bottrop | 120,008 |
| 63 | Pforzheim | 118,460 |
| 64 | Salzgitter | 117,776 |
| 65 | Schwerin | 116,876 |
| 66 | Offenbach am Main | 116,460 |
| 67 | Ulm Universitätsstadt | 115,379 |
| 68 | Zwickau | 112,646 |
| 69 | Ingolstadt | 111,626 |
| 70 | Koblenz | 109,292 |
| 71 | Fürth | 108,011 |
| 72 | Kaiserslautern | 101,970 |
| 73 | Jena | 101,724 |
| 74 | Erlangen | 101,372 |
| 75 | Trier | 99,379 |
| 76 | Dessau | 92,030 |
| 77 | Wilhelmshaven | 90,944 |
| 78 | Brandenburg an der Havel | 87,713 |
| 79 | Flensburg | 87,642 |
| 80 | Neumünster | 82,030 |
| 81 | Neubrandenburg | 81,786 |
| 82 | Frankfurt (Oder) | 81,633 |
| 83 | Worms | 79,737 |
| 84 | Delmenhorst | 78,079 |
| 85 | Plauen | 73,318 |
| 86 | Bayreuth | 72,692 |
| 87 | Bamberg | 69,901 |
| 88 | Görlitz | 68,773 |
| 89 | Stralsund | 66,944 |
| 90 | Aschaffenburg | 66,339 |
| 91 | Weimar | 62,257 |
| 92 | Greifswald | 61,688 |
| 93 | Kempten (Allgäu) | 61,494 |
| 94 | Hoyerswerda | 61,441 |
| 95 | Landshut | 59,257 |
| 96 | Rosenheim | 58,704 |
| 97 | Schweinfurt | 55,598 |
| 98 | Suhl | 53,986 |
| 99 | Neustadt an der Weinstraße | 53,828 |
| 100 | Baden-Baden | 52,677 |
| 101 | Hof | 52,628 |
| 102 | Emden | 51,653 |
| 103 | Passau | 51,035 |
| 104 | Wismar | 50,870 |
| 105 | Speyer | 49,575 |
| 106 | Pirmasens | 48,562 |
| 107 | Frankenthal (Pfalz) | 47,946 |
| 108 | Eisenach | 45,642 |
| 109 | Amberg | 44,177 |
| 110 | Straubing | 44,022 |
| 111 | Coburg | 43,948 |
| 112 | Weiden i.d.OPf. | 43,171 |
| 113 | Kaufbeuren | 42,694 |
| 114 | Memmingen | 40,492 |
| 115 | Ansbach | 39,638 |
| 116 | Landau in der Pfalz | 39,632 |
| 117 | Schwabach | 37,564 |
| 118 | Zweibrücken | 36,039 |

Source: ARBEITSKREIS VOLKSWIRTSCHAFTLICHE GESAMTRECHNUNG DER LÄNDER (2007b).

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ISSN 1865-2441 (PRINT)
ISSN 1865-7133 (ONLINE)

ISBN 978-3-940369-46-8 (PRINT)
ISBN 978-3-940369-47-5 (ONLINE)