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# Biased Survival Expectations and Behaviours: Does Domain Specific Information Matter? 


#### Abstract

We study biased survival expectations across two domains and examine whether such biased expectations influence health and financial behaviors. Combining individual-level longitudinal data, retrospective, and end of life data from several European countries for more than a decade, we estimate time-varying individual level bias in 'survival expectations' (BSE) at the individual level and compare it to biased 'meteorological expectations' (BME). We exploit variation in an individual's family history (parental age at death) to estimate the effect of BSE on health and financial behaviors and compare it to BME, and other tests to discuss whether the effect of BSE results from the effect of private information. We find that BSE increases the probability of adopting less risky behaviors and financial behaviors. We estimate that a one standard deviation increase in BSE reduces the average probability of smoking by $48 \%$ and holding retirement accounts by $69 \%$. In contrast, BME barely affects healthy behaviors, and is only associated with a change in some financial behaviors.


JEL-Codes: I180, D140, G220.
Keywords: biased expectations, survival expectations, meteorological expectations, longevity optimism, private information, health behaviour, financial behaviour.

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[^1]
## 1. Introduction

Rational expectations models (e.g., rational addiction models) are grounded on the assumption that individuals accurately form their survival expectations (Yaari, 1965). Consistently, some research has documented that subjective survival expectations are on average consistent with life table probabilities (Hurd and McGarry 1995, Hurd and McGarry 2002 and Palloni and Novak, 2016). However, already some studies show that aggregate life table realizations do not account for individual-specific heterogeneity (Gan et al., 2004) ${ }^{1}$. The availability of end-of-life data makes it possible to compare more precisely individual level objective and subjective survival expectation evidence ${ }^{2}$. Examining subjective longevity expectations over an individual's life cycle, some studies find evidence of an underestimation (overestimation) of subjective survival at younger (older) age (Elder, 2007, Hurd et al., 2009), suggesting evidence of biased survival expectations. Do such biased expectations affect behaviour? Does the information domain matter?

We study individual-level measures of biased expectations across two domains, namely one's own longevity and the (predicted) weather. More specifically, we measure biased survival expectations (BSE) by comparing for each individual interviewed in end-of-life exit interviews in Europe, their actual survival realizations (and predicted survival for those alive) with their expectations. Similarly, we compare BSE to 'biased meteorological expectations' (BME), that is, how individuals' weather predictions compare to weather realisations.

[^2]Next, upon documenting evidence of biased expectations, we examine the impact of biased expectations on both their survival and the weather on health and financial behaviours. More specifically, we examine the hypothesis of private information driving such effects by exploiting the variation in an individual’s family longevity history (parental age at death) to identify the causal effect of biased expectations on several health (e.g., engaging in preventive actions) and financial behaviours (e.g., saving more). The overestimation of one's subjective survival can be explained by an individual's private information about their objective survival probability, which can come from knowledge of technological or medical advances, or individual-specific genetic survival information ${ }^{3}$. Another test of the role of private information lies in examining whether individual biased expectations vary by domain that differ in the extent of individual control (e.g., limited control of the weather).

We add to the literature in several ways. First, we take advantage of unique data from an end-of-life questionnaire of the Survey of Health Ageing and Retirement in Europe (SHARE) and its retrospective wave (SHARELIFE), which allows us to precisely estimate individual level survival expectations which we then link to the respondent's actual observed survival ${ }^{4}$. Second, we study BSE using longitudinal data from several European countries, which exhibit a large cultural variation in behavioural reference points in the formation of expectations compared to studies using United States data. Third, we differentiate biased expectation in two domains, such as one own survival (or health), and the wealthier (or methodological expectations). Finally, we extend the previous literature ${ }^{5}$ by providing causal

[^3]estimates of biased expectations to specific financial and health behaviours ${ }^{6}$. Hence, we add to the still growing literature on biased expectations and specifically on the role of private information (Kim et al, 2017).

The next section describes this paper relates to the existing literature. Section three describes the data, and the empirical strategy is reported in section four. We report the main results in section five. Next, section six reports a battery of robustness check, and section seven describes other potential pathways. A final section concludes.

## 2. Related Literature

Biased Survival Expectations (BSE). Subjective survival expectations are central to long term decisions regarding saving and consumption (Levhari and Mirman, 1977; Hamersmesh, 1985; Hurd, 1989), retirement and employment (Hurd et al, 2004; Cocco and Gomes, 2011), as well as risky health behaviours (Khwaja et al, 2007). However, the evidence of bias in expectation formation mainly relies on the comparison between cross-sectional expectation data and survival life tables. Consistently, Post and Hanewald (2013) show that the dispersion of subjective survival expectations is more in line with the dispersion of human mortality tables. However, life table models are affected by selection, individual heterogeneity as well as potential cohort effects and non-response of those with higher mortality risk (Bissonette et al, 2017). In contrast, exit interviews are less affected by selection and can precisely identify survival at the individual level.

[^4]Cognitive Biases and Optimism. It is unclear from the literature whether BSE are the result of cognitive biases or the result of other explanations. Individuals differ in their ability to judge their own mortality risk (Lichtenstein et al, 1978), due to differences in personality and emotions. Younger respondents tend, on average, to underestimate their objective survival chances, and the opposite is true for older people (Gronek et al, 2017). Although Steffen (2009) found that both men and women underestimate their life expectancy compared to actuarial life expectancy, Teppa and Lafourcade (2013) estimate that women exhibit a systematic lower subjective survival probability relative to actuarial survival probabilities. Mirowski (1999) argues that surviving at older ages may be a source of optimism if the individual believes that is "increasingly remarkable". Hence, optimism can stand out as a potential explanation for biased survival expectations. Biased longevity expectations should be distinguished from more general forms of optimism, as we discuss later in the paper.

In estimating survival probabilities, there are important reporting biases to account for such the rounding of the data and the presence of focal points (e.g., 0.5) as discussed in Bruine et al., (2000). A nontrivial question is how to estimate survival expectations by age from expectation questions of survival of an additional year of life, which is dealt with examining a subjective hazard function under some distributional assumptions (see Bissonette et al., 2017). Viscusi and Hakes (2003) estimate subjective survival probability and find that they capture a generalised assessment of an individual health status different from a subjective probability as such. Consistently, age, gender and socio-economic status are systematically found to influence subjective survival expectations (Hamermesh, 1985; Hurd and McGarry, 1995; Khwaja et al, 2007).

Behaviour and BSE. Although some studies document the presence of discrepancies between perceived and realized survival, there is limited, or no consensus on the effect of BSE on actual behaviour. Gronek et al, (2017) find no evidence that biased survival expectations affect savings behaviours. Previous studies have examined the effect of survival expectations on retirement (O’Donnell et al., 2006; van der Klaauw and Wolpin, 2008), demand for annuities (Schulze and Post, 2010; Teppa and Lafourcade, 2013), portfolio allocation (Kézdi and Willis, 2011), education (Arcidiacono et al., 2012), migration (McKenzie et al, 2013), savings (Bloom et al., 2006) and smoking behaviour (Balia, 2011). However, most such studies do not account for the endogeneity of survival expectation formation on actual survival. This paper exploits the evidence on parental survival to account for such endogeneity drawing from an instrumental variable strategy.

## 3. The Data

We use data from SHARE (Survey of Health, Ageing and Retirement in Europe) ${ }^{7}$. This survey is the European equivalent of the Health and Retirement Survey. We use data from the "Main Questionnaire" of waves 1 (2004-05), 2 (2006-07), 3 (2009), 4 (2011), 5 (2011) and 6 (2015) and the "End of life" module for waves 2 (2006-07), 3 (2009), 4 (2011-13), 5 (2013) and 6 (2015). We also use and the retrospective SHARELIFE data in our robustness checks. Our final samples are obtained by following several steps (see Table B1 from Appendix B). Using the initial samples for waves 1 to $6(\mathrm{~N}=260,244)$, we have selected those countries which appear in all waves (i.e, Austria, Belgium Denmark, France, Germany, Italy, Spain, Sweden, and Switzerland; $\mathrm{N}=160,388$ ). Second, we select observations with non-missing values for

[^5]calibrated sampling weights ${ }^{8}(\mathrm{~N}=156,320)$. Third, we merge two consecutive waves of SHARE, in the case of survivors two consecutive "Main Questionnaires" of SHARE, whereas for the deceased we merge the "Main Questionnaire" of one wave with the "End-of-life Questionnaire" ${ }^{9}$ of the next wave ( $\mathrm{N}=118,025$ ). We identify there are 29,376 non-follow-up respondents, that is, respondents who have only participated in one wave. Next, we compare the characteristics of the non-follow-up respondents with those of the survivors and deceased subsamples. Finally, we distinguish those respondents who have answered the life expectancy module from those who have not. ${ }^{10}$ Retention rates ${ }^{11}$ increase remarkably over time in all countries resulting in very high overall panel stability after several waves (Table B2 from Appendix B).

Subjective survival. The "Main Questionnaire" contains an "Expectations" module which starts with a warm-up question asking, "What do you think the chances are that it will be sunny tomorrow?". This should help respondents feel at ease with the numerical scale used in the whole set of questions in this module. To make understanding easier these questions are accompanied by cards, with a numerical sequence of probability: from 0 , meaning means absolutely no chance, to a value of 100 , meaning absolutely certainty. Furthermore, respondents are asked to state their Subjective Survival Probability (SSP) on a scale from 0 to 100 using the following question: "What are the chances that you will live to be age [T] or more?". The target T takes the values $\{75,80,85,90, \ldots, 120\}$ depending on the age of the respondent. Respondents aged under 65 at the time of the interview are presented a target of

[^6]75 , while those aged between 65 and 69 have a target of 80 , and so on ( $T=85$ for $70-74$ years, $\mathrm{T}=90$ for $75-79$ years, $\mathrm{T}=95$ for $80-84$ years, $\mathrm{T}=100$ for $85-89$ years, $\mathrm{T}=105$ for $95-100, \mathrm{~T}=110$ for 100-104 years and $\mathrm{T}=120$ for 105 years and older).

Biased Survival Expectations (BSE). Figure B1 shows the evolution of SSP across the six waves. We find no significant differences and a parsimonious reduction of survival probabilities as individuals becomes older. Given the significant reduction in sample size, we compare the density function of the survival subjective probability between follow-up respondents and non-follow up respondents. Regarding the validity of SHARE information, Schulz and Doblhammer (2011) compared age-specific death rates for all SHARE countries of wave 1 with age-specific death rates from the Human Mortality Database. They found that empirical death rates mostly ranged within the confidence interval up to age 65, but from age 65 onwards, they were predominantly below the lower bound of the confidence interval ${ }^{12}$. Consequently, older respondents institutionalized shortly before death may be coded as attrition cases, instead of deceased.

Kernel density functions are shown in Figure B2 (Appendix B) and results of the test of equality of distributions are reported in Table B3. When examining all waves pooled or paired waves, we identify similar density functions, which leads us to reject the null of equality of distributions ${ }^{13}$. Table B4 displays the mean survival subjective probabilities by country and target (T) using the final sample ( $\mathrm{N}=69,900$ ). There is large cross-country heterogeneity in survival expectations across European countries. Such evidence is confirmed by F tests comparing the highest and lowest means across countries. Danish respondents show the largest survival expectations for targets 75,80 and 85 , whereas Italians have the highest survival

[^7]expectations for the oldest cohorts ${ }^{14}$. In contrast, Belgians report the lowest expectations for the youngest cohort, Germans are the least optimistic for targets 80,85 and 90 , and finally, Swedish aged 90 and older are the most pessimistic of all the countries considered ${ }^{15}$.

## 4. Empirical Strategy

### 4.1. Main specification

Using the Weibull specification (commented in section 4.1; see Appendix C for the explanation of the duration model), we estimate the predicted survival probabilities, which we denote as "objective survival probabilities" $\left(O S P_{i c t}\right)$. Accordingly, we define our longevity optimism (or "biased survival expectations") indicator as the difference between subjective and objective survival expectations as follows:

$$
\begin{equation*}
B S E_{i c t}=S S P_{i c t}-O S P_{i c t} \tag{1}
\end{equation*}
$$

Next, we measure the effect of longevity optimism ( $B S E_{i c t}$ ) for an individual $i$ living in country $c$ at time (wave) $t$; on several behaviours' $\left(D_{i c t}\right)$ related to health and financial domains (health, lifestyle, savings, financial investments, and risk taking; see Table A3), as follows:

$$
\begin{equation*}
D_{i c t}=\alpha_{0}+\alpha_{1} B S E_{i c t}+\alpha_{2} X_{i c t}+T_{t}+C_{c}+\epsilon_{i c t} \tag{2}
\end{equation*}
$$

where $D_{i t c}$ refers to the behaviour indicator considered for individual $i$ living in country $c$ at time (wave) $t ; X_{i t c}$ is a vector of explanatory variables ${ }^{16}, T_{t}$ and $C_{c}$ denotes wave and country

[^8]fixed effects and $\epsilon_{i}$ is an error term. We include three groups of explanatory variables: (i) sociodemographic characteristics, (ii) economic controls and (iii) health-related controls. Additionally, we include country fixed effects and wave fixed effects. The definition of each is shown in Table A2.

### 4.2. Endogeneity of Biased Survival Expectations on Behaviour

Given the potential for endogeneity, we follow an instrumental variable approach. In the first stage, we estimate (3) by BSE:

$$
\begin{equation*}
B S E_{i c t}=\beta_{0}+\beta_{1} M A D_{i c t}+\beta_{2} F A D_{i c t}+\beta_{3} M L_{i c t}+\beta_{4} F L_{i c t}+\beta_{5} X_{i c t}+T_{t}+C_{c}+\varepsilon_{i c t} \tag{3}
\end{equation*}
$$

where $M A D_{i t c}$ and $F A D_{i t c}$ represent mother's and father's age at death, respectively, $M L_{i c t}$ and $F L_{i c t}$ are binary variables that take the value 1 if the mother (father) is alive at the time of the survey binary, and 0 otherwise. A valid instrument should be uncorrelated with the error term of (2) but correlated with $B S E_{i c t}$. Hurd and McGarry (2002) and Dormont et al, (2014) have shown that death of a parent is associated with a reduction in subjective survival expectations. Additionally, genetic factors or parental ill health are a form of private information that predicts their descendants' health status, and thus, affect their survival expectations. Therefore, the validity of our instruments relies on the fact that parent's living status affects children's behaviours only through the channel of their SSP. We have selected four instruments: father's age at death, mother's age at death, whether the father and the mother are alive. Here, we exploit link between parents and children objective longevity ${ }^{17}$.

[^9]In measuring the effect of parent's age of death we have distinguished between individuals those whose parents have already passed away at the time of the survey from those who are still alive. In the former case, we have used the reported age of death. In the latter, we draw on the predicted age of parental death using multiple imputation (Rubin, 1987). A detailed explanation of the imputation process is reported on Appendix B. The density function and descriptive statistics of deceased parents (reported age of decease in SHARE) and living parents (imputed age of decease) are shown on Figure B3 and Table B7 ${ }^{18}$. Sections 6.3 and 6.4 reports robustness checks testing for the validity of our instruments. Finally, the predicted values of $B S E_{i c t}\left(\widehat{B S E}_{i c t}\right)$ are used instead of the original values of $B S E_{i c t}$ in equation (1).

$$
\begin{equation*}
I N D_{i c t}=\gamma_{0}+\gamma_{1} \widehat{B S E}_{i c t}+\gamma_{2} X_{i c t}+T_{t}+C_{c}+\omega_{i c t} \tag{4}
\end{equation*}
$$

For each indicator ( $I N D_{i c t}$ ) we have estimated four different models, gradually introducing new explanatory variables ( $X_{i c t}$ ). Different specification differs depending on sociodemographic controls ${ }^{19}$ country fixed effects and wave fixed effects. Next, consider controls for the number of days spent in hospital in the previous year (due to surgery or medical tests), the number of days stayed at hospital during last year (mental health), the number of days stayed at other institutions during last year and the number of visits to a general practitioner during the previous year ${ }^{20}$.For each indicator and each model specification we provide three test statistics for the IV estimation (see Tables B13 to B16) ${ }^{21}$.

[^10]
### 4.3. Biased Meteorological Expectations

To identify whether individual biased expectations vary by domain, we have constructed an indicator that compares the subjective probability of a sunny day the following day after the interview with the meteorological probability of a sunny day (using weather data). Gaisma website ${ }^{22}$ provides sunrise and sunset, and dusk times for thousands of locations all over the world. For each month and region, we have computed average daylight minutes $\left(D M D_{m, r}\right)$ as the difference between sunrise and sunset times (see Table B9). The World Meteorological Organization ${ }^{23}$ provides the mean number of sunshine hours per month, year, and region. First, we have defined the number of sunshine minutes per month $m$ of year $y$ in region $r\left(S M M_{m, y, r}\right)$, where region $r$ is the equivalent of NUTS-2 of SHARE. Second, we have computed the average sunshine minutes per day in month $m$ of year $y$ in region $r\left(S M D_{m, y, r}\right)$ as the ratio between sunshine minutes per month and days per month for each year (to consider leap years) (see Table B10).

$$
\begin{equation*}
S M D_{m, y, r}=\frac{\text { SMM }_{m, y, r}}{\text { Days per month }} \text { y } \tag{5}
\end{equation*}
$$

The meteorological probability of a sunny day in month $m$ of year $y$ in region $r$ ( $M e t S D_{m, y, r}$ ) is defined as the ratio between sunlight minutes per day $\left(S M D_{m, y, r}\right)$ and daylight minutes per day $\left(D M D_{m, r}\right)$ (see Table B11):

$$
\begin{equation*}
M e t S D_{m, y, r}=\frac{S M D_{m, y, r}}{D M D_{m, r}} * 100 \tag{6}
\end{equation*}
$$

$M e t S D_{m, y, r}$ is regressed over country and regional fixed effects and the meteorological predicted probability of sunny day $\left(\widehat{M e t S} D_{m, y, r}\right)$ is obtained.

[^11]The expectations' module of SHARE includes the following question: "what do you think the chances are that it will be sunny tomorrow? For example, '90' would mean a 90 per cent chance of sunny weather. You can say any number from 0 to 100". We use the answer to this question to define reported probability of a sunny day by individual $i$ interviewed in month $m$ of year $y$ in region $r\left(\operatorname{Rep}_{i, m, y, r}\right)^{24}$.
"Meteorological expectations" of individual $i$ interviewed in month $m$ of year $y$ in region $r\left(B M E_{i, m, y, r}\right)$ is the difference between reported probability of a sunny day and meteorological predicted probability of a sunny day:

$$
\begin{equation*}
B M E_{i, m, y, r}=\operatorname{Rep} S D_{i, m, y, r}-\widehat{M e t S} D_{m, y, r} \tag{7}
\end{equation*}
$$

Table B12 compares both indicators of biased expectations. Our estimates suggest that: (about 56\% individuals exhibit both positively biased survival and meteorological expectations (i.e, $\mathrm{BSE}>0$ and $\mathrm{BME}>0$ ) and about $8 \%$ have negatively biased expectations for both domains. The share of respondents exhibiting positively biased expectation on both domains is 9.2 pp is higher for men as compared to women, differ by age and are, 3.5 pp higher for survivors as compared to deceased. Finally, we find that Austria and Denmark exhibit the highest share of individuals with positively biased expectations, whereas Belgium, France and Germany exhibit the highest of negatively biased expectations (i.e., with the highest percentage for $\mathrm{BSE}<0$ and $B M E<0$ ). Finally, we estimate OLS regressions for the same set of indicators used in equation (4), but replacing it by expectations based on meteorological beliefs ( $B M E_{i, m, y, r}$ ):

$$
\begin{equation*}
I N D_{i c t}=\delta_{0}+\delta_{1} M E_{i, m, y, r}+\delta_{2} X_{i c t}+T_{t}+C_{c}+\varsigma_{i c t} \tag{8}
\end{equation*}
$$

where $X_{i c t}$ is the same set of control variables used in (4) and we proceed in a similar fashion, adding them gradually; $T_{t}$ and $C_{c}$ are fixed and country fixed effects and $\varsigma_{i c t}$ is an error term.

[^12]
## 5. Results

### 5.1. Biased Subjective Survival

We begin with estimating BSE among survivors and the individual characteristics predicted to make an influence. Table B8 reports estimates of the discrete-time hazard model and identifies a long list of unhealthy lifestyles that correlate with objective survival across waves. Estimates suggest that the hazard rate (1.429) increases over time but follows a decreasing pattern. At each survival time, men exhibit a higher hazard ratio (+88.4\%) hazard ratio than women. Compared to individuals who have never smoked, the hazard rate is $52.4 \%$ higher among current smokers ${ }^{25}$ and $36.6 \%$ higher for past smokers. Consistently, we estimate a higher hazard rate among respondents that fail to perform any moderate nor vigorous physical activity experience larger hazard rate, and those who drink every day or almost every day (23.4\%). Being diagnosed with cancer ${ }^{26}$ and Alzheimer’s increases mortality risk by $156.6 \%$ and $58.5 \%$, respectively, and we document that similarly, those who have suffered a stroke exhibit a 54.3\% higher hazard ratio. Individuals feeling depressed exhibit an 18.1\% higher hazard ratio which confirms previous evidence regarding the relationship between depression and mortality (Mykletun et al., 2009). Finally, we document a negative association between the survival hazard rate and educational attainment ${ }^{27}$.

Figure 1 depicts the subjective survival probability (SSP) and objective survival probability (OSP) for both deceased, and survivors, across individual's ages. Consistently with the hypothesis of biased survival expectations, we find evidence of significant subjective

[^13]biased survival expectations, both among survivors and the deceased. However, such bias widens among the deceased over time whilst it remains relatively stable among survivors.

## [Insert Figure 1 about here]

Table 1 examines age differences in BSE between deceased and survivors. For each age cohort, the column SSP indicates the average subjective survival probability of reaching target T (between 0 and 100) and the column OSP shows the average objective survival probability (between 0 and 100) obtained from the duration model with the Weibull specification ${ }^{28}$. We find that BSE does not vary across the age cohorts 65-69 and 70-74. The last two columns in Table 1 displays the percentage of individuals revealing both a positive and negative bias. More specifically, we define an individual as exhibiting positive BSE if SSP is higher than the predicted OSP, whereas we classify the reverse as negatively biased. Although the percentage of individuals with biased survival expectations among the survivors is stable it peaks at age 95. In contrast, among the deceased, biased survival expectations peak at age 80. One explanation for the overestimation of survival expectations for the older cohort is due to the detrimental effect of ageing on cognitive abilities needed to estimate mortality rates ${ }^{29}$ ( Elder, 2007).

## [Insert Table 1 about here]

[^14]When we estimate BSE by gender (in Table C2 in the appendix) we find that consistently with prior research, men are more likely to exhibit biased survival expectations than women for both deceased and survivors ${ }^{30}$, though the bias is higher both in the deceased subsample, and among women ( 10.2 pp ). This implies that although women tend to reveal more negatively biased expectations, deceased women exhibit more positively biased expectations than surviving women ${ }^{31}$. Similarly, when we examine differences across countries (Table C3) ${ }^{32}$, we find that two thirds of the population exhibit a positive biased expectation and one third a negative biased expectation. Among the sample of diseased, Italians are on average those that reveal the highest positive bias whilst the Swedish reveal highest share of negatively biased. In contrast, among survivors Belgians, Germans and French respondents exhibit the highest share of negatively biased survival and Danish and Spanish the highest share of positive bias survival.

### 5.2. Explaining Biased Survival Expectations

Table D1 examines covariates explaining BSE which make up the first stage instrumental variable regression ${ }^{33}$. The four instruments considered are significant and exhibit the expected signs. They suggest that the effect of age of death and living status is higher for fathers than for mothers. On average, each year of maternal (paternal) age at death raises BSE by 0.0012 points ( 0.0017 points). Individuals with a living mother (or father) exhibit a higher BSE by 0.0114 points ( 0.0248 points). Men are significantly more positively biased than women (12.42 points), and each year of life increases BSE by 7.62 points. Living in Denmark and Spain increases the probability of positive biased survival by 16.5 and 15.1 pp respectively,

[^15]as opposed to living in Germany, where expectations only increase by 2.2 pp . Consistently, those who are coded as having stayed at the hospital are more likely to show negatively biased expectations. On the contrary, stays at other institutions due to convalescence or rehabilitation are associated with only a slight increase in negatively biased expectations.

Time invariant fixed effects reduce the BSE by 30.2 points between waves 2 and 4, which corresponds to the period 2007-2011. The influence of education is U-shaped, revealing that individuals with no education, those with primary education and those with post-secondary tertiary education are significantly more positively biased than those with lower secondary education. Consistently with previous research (Hurd and McGarry, 2002), we find that income plays a larger influence than wealth, and high-income people are less pessimistic. Each 1,000 PPP units of adjusted income (wealth) increase BSE by 66.4 pp (1.2pp).

Figure B4 displays the predicted subjective longevity by respondent's age, gender and parents' living status. When both parents are alive, BSE increases with respondent's age (e.g., BSE for men (women) aged 70-74 years old is eight (seven) times higher as compared to BSE for men (women) aged 50-54 years old). In contrast, when both parents have deceased, we find a considerably smaller effect among over 75. Comparing BSE by respondent's gender, we find that biased survival is higher for men regardless of parents’ living status except for the first age cohort (50-54 years). When both parents have deceased, BSE for men describe an inverted Ushape, peaking at 75-79 years, whereas BSE for women continue relatively stable for the three last age cohorts.

### 5.3. Effects of biased survival expectations on health behaviours

Table 2 reports the estimated effect of longevity optimism on a number of outcomes. Table B13 in the appendix reports the effect of a one standard increase BSE and BME. IV estimates suggest that biased survival expectations reduce the probability of unhealthy
lifestyles such as smoking, alcohol intake and sedentary lifestyles. This evidence is consistent with the hypothesis that individuals who expect to live longer adjust their behaviours accordingly. Estimates suggest that one standard deviation increase in BSE decreases (increase) the probability of smoking (never having smoked) by 13pp (24.9pp), which entail an average decrease (increase) of $48 \%(51 \%)^{34}$. Similarly, a one standard deviation increase in BSE reduce (increase) the probability of daily alcohol consumption (not having drunk alcohol during the last 3 months or less than once per month) in $19.9 \mathrm{pp}(28.8 \mathrm{pp})^{35}$. These results are consistent with previous evidence that finds that binge and heavy binge drinkers think that their driving ability and ability to tolerate alcohol is better than average (Sloan et al. 2013). However, such positive bias does not explain why some adults consume large amounts of alcohol. We estimate that a one standard deviation increase in BSE, reduces sedentary lifestyles, raising the probability of moderate or vigorous physical activity in 24.5pp.

We find that both positive and negative biased subjective expectations (BSE) exert comparable effects on smoking behaviour. However, we identify heterogeneous effects on other behaviours. When we examine the effects of meteorological expectations (ME), we observe a similar sign and consistent effect, but its magnitude, except for "sedentary lifestyles", is almost negligible.

## [Insert Table 2 about here]

### 5.4. Effects on health behavioural outcomes

[^16]Next, when we turn into examining anthropometric and health outcomes (Table 3) ${ }^{36}$, we document that a positive BSE increases the likelihood of a normal weight and increases the likelihood of obesity. A one standard deviation increases in BSE increases (decreases) the probability of being normal weight (obese) by 17.9pp (27.8pp), which implies an increase (decrease) by $47.5 \%$ (46.7\%) with respect to the mean probability of being normal weight (obese). These results confirm previous evidence that individuals with BSE generally have better physical health.

### 5.5. Effects on cognitive abilities

Next, we find that individuals with biased survival expectations can recall more words, both in the first and the second recall. One standard deviation increase in an individual's BSE increases the number of words recalled ( 0.52 in the first recall, 0.55 in the second one). Furthermore, we find that BSE increases maximum grip strength test scores. A one standard deviation increase in an individual's BSE leads to an increase in the maximum grip strength test by 1.47 kg , or an average increase in $4.3 \%$ with respect to the average score. Consistently, negative BSE deliver a consistent effect with an opposite sign, but effect sizes are smaller in magnitude. However, when we look at effect sizes, we find that positively biased expectations ( $\mathrm{BSE}>0$ ) produce a larger effect than negatively biased expectations ( $\mathrm{BSE}<0$ ). In contrast, the effect for meteorological optimism albeit consistent, is more modest.

## [Insert Table 3 about here]

### 5.6. Effects on financial behaviour ${ }^{37}$

[^17]We then estimate the effect of biased survival expectations (BSE) on the probability of owning financial assets and including cash and increases household wealth. One standard deviation change in BSE increases the probability of ownership of bonds by 2pp, mutual funds by 3.8pp and individual retirement accounts by $5 \mathrm{pp}^{38}$. Consistently, it decreases the probability of having bank accounts by 17.4 pp , investing in stocks by 2.3 pp , having a mortgage by 5.4 pp and having other debts by $5.3 \mathrm{pp}^{39}$. Consistently, negatively biased survival expectations $(\mathrm{BSE}<0)$ explain some financial investments ${ }^{40}$. These results point to an intense deaccumulation of assets, and paradoxically, at the same time a higher concern for the wellbeing of the family (purchase of life insurance).

Compared to health behaviours, biased meteorological expectations exert a more modest effect on a few financial behaviours. If heterogeneity in BME is mostly due to private information, one might expect the absolute difference to correlate with behavior or lifestyle. For instance, people who work outside the household, or follow the news may get more accurate weather estimates than the authors' benchmark. However, we find that negative and positive ME seem to have opposite effects. Indeed, we find that a one standard deviation changes in meteorological optimism (BME) increases (decrease) investments in bonds (life insurance) by $13.5 \%$ (22.5\%) with respect to its average probabilities. It also increases the probability of having a mortgage by $36.92 \%$. In comparison, the effect of one standard deviation increases in negatively biased meteorological expectations ( $\mathrm{BME}<0$ ) is more modest (10.5\% and $14.5 \%$, respectively).

[^18]
## [Insert Table 4 about here]

### 5.7. Effects on financial risk-taking survival ${ }^{41}$

Finally, Table 5 reports the effect of BSE on above average financial risk taking, and documents that one standard deviation increase in negatively biased survival (BSE<0) increases the uptake of prudent financial behaviours. Finally, we on average find that ethe effect of positively biased survival expectations ( $\mathrm{BSE}>0$ ) is twice that of negatively biased expectations ( $\mathrm{BSE}<0$ ). The effects for MO are in line with those of BSE but are significantly smaller.

## [Insert Table 5 about here]

## 6. Robustness checks

### 6.1. Focal Responses

One potential explanation of our results could be the presence of focal responses. Focal responses refer to responses clustered around certain 'focal points' of the distribution (Hurd and McGarry, 2002). From the histograms (Figure D1) it becomes clear that this is indeed the case in our study. For example, $17.2 \%$ of the deceased reported a null probability, which refers to $20.8 \%$ of the survivors and $18.9 \%$ of the deceased sample reported a survival probability of $50 \%$, and $17 \%$ of the survivors reported a probability of $100 \%$. Some authors suggest that these responses may constitute a signal of inability for answering probabilistic questions (Bruine de Bruine et al.,2002), while others refer to these responses also contain valuable information (Manski and Molinari, 2010). Bruine de Bruin et al. (2002) coined the term "epistemic uncertainty" to describe the behaviour of respondents who reported a 50 per cent probability.

[^19]They consider that in this case where the respondent does not have any expectations at all, the response would be equivalent to "don't know".

Some authors claim that the likelihood of providing focal points is positively associated with lower cognitive performance. Perry (2005) found that individuals answering focal responses of zero and one were on average less educated, held fewer assets and had lower income than the rest of the sample. On the contrary, those respondents reporting a 50 percent chance of surviving compared to the rest of the sample. Likewise, Hurd et al. (1998) and Kleinjans and van Soest (2014) observed that more educated individuals were less prone to reveal focal points. Table D3 reports cognitive and economic related variables for focal and non-focal respondents and the corresponding comments are shown below the table ${ }^{42}$.

As a robustness check, we re-estimate the model excluding focal responses. Table B1 shows the characteristics of the sample for survivors and deceased after excluding focal responses ( 0,50 and 100$)^{43}$. The estimated coefficients and hazard ratios for the preferred Weibull specification are reported in Table B8 ( ${ }^{\text {rd }}$ column). Our estimates suggest that the magnitude and significance is very similar to the original sample. Figure D2 compares the density function for the "biased survival" indicator including and excluding focal points. Again, differences between both density functions are almost negligible, which suggests that the inclusion of focal responses does not distort the conclusions of the paper (as has been also observed by Kleinjans and van Soest, 2014).

### 6.2. Respondents' understanding of probabilities

[^20]A prerequisite for the consistent interpretation of subjective survival expectations is that respondents can answer probabilistic questions (Manski, 2004). Hence, as an additional robustness check, we have excluded those individuals who provided an erroneous response to the following probabilistic question: "If the chance of getting a disease is 10 per cent, how many people out of 1000 (one thousand) would be expected to get the disease?" ${ }^{44}$. Again, compared to the baseline coefficients and hazard ratios, estimates resulting from the reestimated model are of similar magnitude. Figure D3 shows the kernel density function of the "biased survival" indicator for both the total sample and the reduced sample after applying the exclusion criteria. Importantly, the shape of both density functions does not suggest evidence of bias.

### 6.3. Use of (private) medical information from SHARELIFE

Another concern is that the imputed parental age at death refers to parents who are alive at the time of the survey, but parental age at death is not time varying. Hence, we use retrospective data from parental vital status available in SHARE Wave 3, known as SHARELIFE ${ }^{45}$. We can test for the effects of private information with regards to one's own health consist with other work (Viscusi and Hakes, 2003). Hence, we have re-estimated our models using two additional instruments: blood pressure being checked the previous year and having his/her blood tested ${ }^{46}$ both proxies for potential private health information.

[^21]Consistently, we find that the magnitude and significance of the estimated coefficients compares to that of the entire sample ${ }^{47}$.

Table D1 (4 $4^{\text {th }}$ column) reports the results of the IV regression for BSE, and we find that the values of the corresponding instruments to parent's age of death and parent's living status compare to those using the entire sample. Having blood pressure checked in the last year reduces BSE in 11.6 and having blood tested in the last year reduces BSE by 8.1pp. However, Table D4 reveals that even after such adjustment, the effect of the BSE confirms our previous conclusions.

Figure D4 compares observed BSE for the total sample, with predicted BSE for the subsample of respondents who also completed SHARELIFE using all instruments (parents' age of decease, blood pressure and blood check) and predicted BSE using only two instruments (blood pressure and blood check; $5^{\text {th }}$ column of Table D1). Although the differences are not large, we find that when all the instruments are used the predicted BSE is closer to the observed BSE.

### 6.4. Instrument validity

In interpreting BSE as the difference between objective and subjective survival expectations, one potential interpretation of our estimates is that private information (unknown to the econometrician) influences expectations above and beyond other more general biases such as optimism. To examine these concerns more closely, Figure D5 compares the observed BSE, the predicted BSE (according to estimations of Table D1) and the predicted BSE taking the average for each survey respondent. Consistently with the hypothesis that BSE reflects partially an effect of unknown private information, we observe that, for both survivors and

[^22]deceased samples, the predicted BSE is smaller but still significant and exhibit a large effects size, in line with previous estimates.

To check the reliability of the instruments, we re-estimate the first-stage regression using lagged instruments. Although the sample size is considerably reduced, Table D5 shows that the instrumental variable estimates ${ }^{48}$, are relevant and significant, and robust. Similarly, to enhance the causal claims, we examine the two bound weak instrument method proposed by Conley et al. (2012) that allows us to retrieve inferences even when the instrumental variables do not satisfy the exogeneity restriction (see Appendix E for explanation of both approaches). Figure E1 shows estimates using as the instrument "father's age of death" (similar results have been obtained for the other instruments; results available upon request). On the left side we show the graphs for the Local-to-Zero approach. The solid line depicts the 2SLS father's age of death effect estimate for the respective indicator (for simplicity, we only show the results for one indicator of each group). The two dash lines depict upper and lower limits of the respective test scores. Overall, the results confirm that even with substantial deviation from the exclusion restriction, the instrument still has a considerable effect over the outcome variable ${ }^{49}$.

Finally, Figure E1 also shows the $95 \%$ confidence bounds of the instrumental variable’s coefficient using father's age of death). Taking as reference the zero-line (red-line), once the confidence bounds include the zero-value, the 2SLS are no longer significant at the $95 \%$ level of significance. On the contrary, if the upper limit crosses the zero-line at a high value of $\delta$, then the 2SLS are robust to possible violations of the exclusion restriction assumption. Similar

[^23]figures have been obtained for the instrument "mother's age of death" (figures available upon request).

### 6.5. Inclusion of parental characteristics in the discrete hazard model

As a final check, we re-estimate the discrete hazard model including as explanatory variables the age at death of the parents and the fact of having a living parent. Table C1 shows that, once again, the Weibull model is the preferred model. Using this specification, Table E2 shows the estimated coefficients and hazard ratios for the initial sample, excluding focal responses, excluding those who answered probabilistic questions erroneously and for the subsample that answered SHARELIFE. with the results compare to those of Table B8. Although living mother and father ins now non-significant, age at death reveals the expected negative effect on the probability of death.

## 7. Mechanisms.

The Nature of Private information. To examine whether BSE convey the effect of the type of respondent's private information, or measurement error we estimate the effect lagged BSE as well as of the death of a brother, sister, or child as other form of private information in addition to one's parents. The results (Table E1 in the appendix) show that the effect of lagged BSE is significant at $1 \%$ and this significance is preserved when the explanatory variables are progressively included, and when the death of the brother, sister or child is added. However, the death of a sibling is not significant at $5 \%$ level, nor is the death of a child exert a significant either. Hence, we conclude that the nature of private information matters, as only parental longevity plays a role in expectation formation.

Finally, we examine the effect of two additional pathways that can be driving the effect, one is the effect of BSE on social contacts, as well as the individual sense of control of one's life, so called locus of control.

Social contacts. Individuals with BSE might engage in more social contacts and tend to decrease in the probability of sedentary lifestyles. We estimate equations (4) and (8) using as the dependent variable a binary indicator that takes the value 1 if the respondent has participated in the last month in voluntary or charity work, gone to a sports or social club, taken part in political community activities, a religious organization or attended educational or training courses, 0 otherwise (see Table A3 for definitions). Table B17 reports the estimated coefficients and Table B18 reports the effect of a one standard increase of BSE and ME. We find that one standard deviation change in BSE increases the probability of having social contacts by 6.08 pp (which represents $38.72 \%$ with respect to the mean probability), which is six times higher than the effect of one standard deviation is positive BSE. Similarly, we find that optimistic ME increase social contacts (5.8 pp or increase by 37\% with respect to the mean probability).

Locus of control. Another potential effects reforms to changes in an individual's sense of control of their life as they age, which we refer to as locus of control as initially proposed by Rotter (1954). We estimate equations (4) and (8) using as dependent variables four binary indicators that take the value 1 if respondent answers that they feel that things are out of control: often, sometimes, rarely, and never. Table B17 reports the estimated coefficients and Table B18 reports the effect of a one standard increase of BSE and ME.

Consistently, we find that BSE is associated with a reduction in the probability of losing of control. A one standard deviation increase in BSE increases the probability of "feeling things are out of control: often" by $152.57 \%$ and increases the probability of "feeling things are out of control: never" by 42.36\%. In contrast, longevity pessimism (BSE $<0$ ) gives rise to a large
increase in the probability of often feeling overwhelmed (194.2\%). In contrast, the effects of BME, are not as robust as those of BSE. These results confirm that BSE may give rise to active coping strategies and the feeling of control of their life (Lopes and Cunhna, 2008).

## 8. Conclusion

We study how biased expectations of one survival and the predicated weather affect health and financial behaviours. We have estimated biased survival expectations (BSE), computed as the difference between actual and subjective survival using individual level longitudinal evidence of individuals in a long list of European countries. Similarly, we estimate a measure of biased meteorological optimism (BMO) comparing as the difference between individual level weather predictions and weather realisations. Next, we draw on an instrumental variable analysis that takes advantage of rich data on parental survival to estimate the effect of BSE on health and financial behaviour, which is compared to the effect of BMO.

We estimate evidence of a bias between subjective and objective survival expectations in both domains, and such BSE influences health and financial behaviours. However, we find that a symmetric effect of positively and negatively biased survival expectations on behaviours ${ }^{50}$. Although we document some consistent variation with BMO, we find significant differences in biased survival across domains which explain different effects on behaviours. BSE influences some health and financial behaviours such as smoking, sedentary lifestyles, purchasing life insurance and saving for old age. However, BMO does not influence health behaviours. One standard deviation change in BSE is associated with an increase in the probability of healthier habits (never having smoked by $51 \%$, not having drunk in the last 3 months by $94.83 \%$ and not having a sedentary lifestyle by $211.5 \%$ ). Similarly, one standard

[^24]deviation change in BSE increases the probability of taking above average financial risks expecting to earn above average returns (83.1\%), investing to a layer extent in fixed income securities, such as corporate bonds by $45.45 \%$ and mutual funds (mostly in bonds) by $52.3 \%$, and individual retirement accounts (69.2\%).

In contrast, BMO increases the probability of owning bonds (13.49\%), mortgage (36.9\%), but decreases the probability of purchasing life insurance (-22.5\%). These results are consistent with evidence that biased expectations are domain specific. A clear policy implication emerges if a more positive biased in expectation formation can be learned during a period of life. The latter might result from private information, as well as optimism though our estimates cannot fully distinguish between the two. If we interpret our results are reflecting optimism (positive expectations about the future) then our estimates are consistent with twin studies showing that optimism during individuals’ impressionable years can influence health behaviours (Plomin, 1992).

Our estimates help explain results from Wang and Sloan (2018) who draw attention to the negative effects arising from the underestimation of the risks associated with non-adherence to prescribed treatment. In their analysis of patients diagnosed with diabetes, they note that many sufferers are not conscientious in their adherence to medical recommendations. Similarly, Picone et al. (2004) document that women with biased survival expectations are more likely to perform breast self-exams and to request Pap smears and mammograms consistently with our estimates suggesting positive effects on "healthy" lifestyle, (i.e., smoking, drinking, physical exercise) of BSE. Hence, we conclude that BSE are the result of a combination of private information (e.g., parental age at death) and cognitive biases such as optimism. The formation of such expectations varies across domains and exerts economically relevant effects in explaining health and financial behaviours.

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## Figures and Tables

Figure 1. Comparison of subjective survival probability (SSP) and objective survival probability (OSP) by age of the respondent


Note: These figures display objective survival expectations (OSP) and subjective survival expectations (SSP) for the subsample of survivors and deceased. The distance between OSP and SSP is what we have defined as "biased survival expectations" in equation (1).
Subjective survival probabilities (SSP) correspond to the answer to the question of SHARE questionnaire "What are the chances that you will live to be age $[T]$ or more?". The target $T$ takes the values $\{75,80,85,90, \ldots, 120\}$ depending on the age of the respondent. Respondents aged under 65 at the time of the interview are presented a target of 75 , while those aged between 65 and 69 have a target of 80 , and so on ( $\mathrm{T}=85$ for 70-74 years, $\mathrm{T}=90$ for $75-79$ years, $\mathrm{T}=95$ for $80-84$ years, $\mathrm{T}=100$ for $85-89$ years, $\mathrm{T}=105$ for $95-100, \mathrm{~T}=110$ for 100-104 years and $\mathrm{T}=120$ for 105 years and older).

Table 1. Comparison of subjective survival probability and objective survival probability by age at baseline

| Age at baseline | N | Survival expectations |  |  |  |  | Meteorological expectations |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | $\begin{aligned} & \hline \text { SSP } \\ & \text { (\%) } \end{aligned}$ | $\begin{aligned} & \hline \text { OSP } \\ & \text { (\%) } \end{aligned}$ | BSE Difference SSP-OSP | $\begin{gathered} \text { BSE>0 } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \text { BSE<0 } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \text { BME>0 } \\ (\%) \end{gathered}$ | $\begin{gathered} \mathrm{BME}>0 \\ (\%) \end{gathered}$ |
| 50-64 | 55,609 | $\begin{gathered} \hline 67.79 \\ (26.94) \end{gathered}$ | $\begin{gathered} 56.22 \\ (25.39) \end{gathered}$ | $\begin{gathered} 11.57 \\ (33.89) \end{gathered}$ | 63.96 | 36.04 | 74.87 | 25.13 |
| 65-69 | 7,065 | $\begin{gathered} 49.43 \\ (31.21) \end{gathered}$ | $\begin{gathered} 20.07 \\ (20.14) \end{gathered}$ | $\begin{gathered} 29.36 \\ (36.78) \end{gathered}$ | 75.64 | 24.36 | 72.77 | 27.23 |
| 70-74 | 4,404 | $\begin{gathered} 40.84 \\ (31.30) \end{gathered}$ | $\begin{gathered} 11.20 \\ (15.61) \end{gathered}$ | $\begin{gathered} 29.64 \\ (34.79) \end{gathered}$ | 76.69 | 23.30 | 75.70 | 24.30 |
| 75-79 | 2,110 | $\begin{gathered} 33.04 \\ (31.97) \end{gathered}$ | $\begin{gathered} 3.84 \\ (8.00) \end{gathered}$ | $\begin{gathered} 29.19 \\ (32.99) \end{gathered}$ | 73.87 | 26.13 | 68.34 | 31.66 |
| 80-84 | 617 | $\begin{gathered} 34.86 \\ (30.90) \end{gathered}$ | $\begin{gathered} 1.41 \\ (4.28) \end{gathered}$ | $\begin{gathered} 33.45 \\ (31.39) \end{gathered}$ | 78.28 | 21.72 | 74.35 | 25.65 |
| 85-94 | 83 | $\begin{gathered} 31.88 \\ (31.09) \end{gathered}$ | $\begin{gathered} 0.76 \\ (2.94) \end{gathered}$ | $\begin{gathered} 31.13 \\ (30.79) \end{gathered}$ | 75.70 | 24.30 | 71.16 | 28.84 |
| 95-100 | 12 | $\begin{gathered} 41.58 \\ (29.51) \end{gathered}$ | $\begin{gathered} 0.00004 \\ (0.00012) \end{gathered}$ | $\begin{gathered} 41.58 \\ (29.51) \end{gathered}$ | 75.83 | 24.17 | 64.84 | 35.16 |
| All | 69,900 | $\begin{gathered} 62.48 \\ (29.87) \\ \hline \end{gathered}$ | $\begin{gathered} 47.00 \\ (29.72) \\ \hline \end{gathered}$ | $\begin{array}{r} 15.48 \\ (34.99) \\ \hline \end{array}$ | 66.53 | 33.47 | 74.49 | 25.51 |


| Age at baseline | N | Survival expectations |  |  |  |  | Meteorological expectations |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  |  |  |  |
|  |  | $\begin{aligned} & \hline \text { SSP } \\ & \text { (\%) } \end{aligned}$ | $\begin{gathered} \hline \text { OSP } \\ (\%) \end{gathered}$ | BSE <br> Difference SSP-OSP | $\begin{gathered} \text { BSE>0 } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{BSE}<0 \\ (\%) \\ \hline \end{gathered}$ | BME $>0$ (\%) | BME $>0$ <br> (\%) |
| 50-64 | 823 | $\begin{gathered} \hline 50.88 \\ (32.22) \end{gathered}$ | $\begin{gathered} \hline 34.64 \\ (28.32) \end{gathered}$ | $\begin{gathered} 16.24 \\ (36.54) \end{gathered}$ | 67.01 | 32.99 | 71.73 | 28.27 |
| 65-69 | 346 | $\begin{gathered} 42.45 \\ (32.26) \end{gathered}$ | $\begin{gathered} 13.28 \\ (18.89) \end{gathered}$ | $\begin{gathered} 29.17 \\ (34.73) \end{gathered}$ | 74.35 | 25.65 | 66.33 | 33.67 |
| 70-74 | 358 | $\begin{gathered} 36.68 \\ (32.00) \end{gathered}$ | $\begin{gathered} 6.57 \\ (12.18) \end{gathered}$ | $\begin{gathered} 30.11 \\ (35.35) \end{gathered}$ | 73.21 | 26.79 | 78.59 | 21.41 |
| 75-79 | 312 | $\begin{gathered} 25.68 \\ (30.29) \end{gathered}$ | $\begin{gathered} 2.05 \\ (5.46) \end{gathered}$ | $\begin{gathered} 23.64 \\ (30.13) \end{gathered}$ | 62.54 | 37.46 | 57.99 | 42.01 |
| 80-84 | 164 | $\begin{gathered} 28.96 \\ (27.94) \end{gathered}$ | $\begin{gathered} 1.57 \\ (3.75) \end{gathered}$ | $\begin{gathered} 27.39 \\ (27.97) \end{gathered}$ | 80.01 | 19.99 | 67.33 | 32.67 |
| 85-94 | 31 | $\begin{gathered} 22.82 \\ (27.32) \end{gathered}$ | $\begin{gathered} 0.62 \\ (3.00) \end{gathered}$ | $\begin{gathered} 22.20 \\ (26.84) \end{gathered}$ | 70.79 | 29.21 | 70.30 | 29.70 |
| 95-100 | 6 | $\begin{gathered} 38.41 \\ (32.36) \end{gathered}$ | $\begin{gathered} 0.00 \\ (0.00) \end{gathered}$ | $\begin{gathered} 38.41 \\ (32.36) \end{gathered}$ | 65.92 | 34.08 | 76.30 | 23.70 |
| All | 2,040 | $\begin{gathered} 40.63 \\ (32.87) \\ \hline \end{gathered}$ | $\begin{gathered} 17.18 \\ (24.54) \\ \hline \end{gathered}$ | $\begin{gathered} 23.44 \\ (34.82) \\ \hline \end{gathered}$ | 70.04 | 29.96 | 70.02 | 29.98 |
| Survivors subsample (by age at baseline) |  |  |  |  |  |  |  |  |
| Age at baseline | N | Survival expectations |  |  |  |  | Meteorological expectations |  |
|  |  | $\begin{aligned} & \hline \text { SSP } \\ & (\%) \end{aligned}$ | $\begin{aligned} & \hline \text { OSP } \\ & \text { (\%) } \end{aligned}$ | BSE Difference SSP-OSP | $\begin{gathered} \mathrm{BSE}>0 \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \text { BSE<0 } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { BME>0 } \\ (\%) \end{gathered}$ | $\begin{gathered} \mathrm{BME}>0 \\ (\%) \end{gathered}$ |
| 50-64 | 54,786 | $\begin{gathered} 68.06 \\ (26.76) \end{gathered}$ | $\begin{gathered} 56.57 \\ (25.19) \end{gathered}$ | $\begin{gathered} 11.49 \\ (33.84) \end{gathered}$ | 63.91 | 36.09 | 74.92 | 25.08 |
| 65-69 | 6,719 | $\begin{gathered} 49.81 \\ (31.12) \end{gathered}$ | $\begin{gathered} 20.44 \\ (20.14) \end{gathered}$ | $\begin{gathered} 29.37 \\ (36.89) \end{gathered}$ | 75.71 | 24.29 | 73.11 | 26.89 |
| 70-74 | 4,046 | $\begin{gathered} 41.28 \\ (31.20) \end{gathered}$ | $\begin{gathered} 11.69 \\ (15.85) \end{gathered}$ | $\begin{gathered} 29.59 \\ (34.73) \end{gathered}$ | 77.06 | 22.94 | 75.39 | 24.61 |
| 75-79 | 1,798 | $\begin{gathered} 34.21 \\ (32.08) \end{gathered}$ | $\begin{gathered} 4.13 \\ (8.30) \end{gathered}$ | $\begin{gathered} 30.08 \\ (33.35) \end{gathered}$ | 75.68 | 24.32 | 69.99 | 30.01 |
| 80-84 | 453 | $\begin{gathered} 37.24 \\ (31.74) \end{gathered}$ | $\begin{gathered} 1.34 \\ (4.48) \end{gathered}$ | $\begin{gathered} 35.90 \\ (32.38) \end{gathered}$ | 77.58 | 22.42 | 77.18 | 22.82 |
| 85-94 | 52 | $\begin{gathered} 35.90 \\ (32.03) \end{gathered}$ | $\begin{gathered} 0.82 \\ (2.94) \end{gathered}$ | $\begin{gathered} 35.08 \\ (31.80) \end{gathered}$ | 77.88 | 22.12 | 71.54 | 28.46 |
| 95-100 | 6 | $\begin{gathered} 47.08 \\ (27.31) \end{gathered}$ | $\begin{gathered} 0.00006 \\ (0.00015) \end{gathered}$ | $\begin{gathered} 47.08 \\ (27.31) \end{gathered}$ | 92.98 | 7.02 | 45.01 | 54.99 |
| All | 67,860 | $\begin{gathered} 63.21 \\ (29.48) \end{gathered}$ | $\begin{gathered} 47.99 \\ (29.35) \\ \hline \end{gathered}$ | $\begin{gathered} 15.21 \\ (34.96) \\ \hline \end{gathered}$ | 63.41 | 36.59 | 74.65 | 25.35 |

Note: SSP: subjective survival probability. Indicates the average subjective survival probability of reaching target age T ( $\mathrm{T}=75$ years if Age $<65$; $\mathrm{T}=80$ years if $65 \leq$ Age $<70$; T= 85 years if $70 \leq$ Age $<75$; T= 90 years if $75 \leq$ Age $<80$; T= 95 years if $80 \leq$ Age $<85$; T= 100 years if $85 \leq$ Age $<94$; T= 105 years if $95 \leq$ Age.OSP: objective survival probability. Indicates the average objective survival probability predicted from duration model for the same target age T. Standard errors in brackets. Using calibrated sampling weights

Table 2. Effect of one standard deviation increase (or decrease) of biased survival expectations (BSE) and biased meteorological expectations (BME) over lifestyle variables

|  | One standard deviation increase of BSE>0 | One standard deviation increase of BSE<0 | One standard deviation increase of BME $>0$ | One standard deviation increase of BME $<0$ |
| :---: | :---: | :---: | :---: | :---: |
| Smokes now Effect in percentual points With respect to the mean probability | $\begin{aligned} & -13.0761 \mathrm{pp} \\ & -48.4122 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & 14.6422 \mathrm{pp} \\ & 54.2103 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.1927 \mathrm{pp} \\ & -0.7134 \% \\ & \hline \end{aligned}$ | $\begin{gathered} 0.1983 \mathrm{pp} \\ 0.7342 \% \\ \hline \end{gathered}$ |
| Has never smoked <br> Effect in percentual points <br> With respect to the mean probability | $\begin{gathered} 24.9441 \mathrm{pp} \\ 51.0941 \% \end{gathered}$ | $\begin{aligned} & -26.5743 \mathrm{pp} \\ & -54.4333 \% \end{aligned}$ | $\begin{gathered} 0.1171 \mathrm{pp} \\ 0.2399 \% \end{gathered}$ | $\begin{gathered} -0.3261 \mathrm{pp} \\ -0.6680 \% \end{gathered}$ |
| Drinks every day or $\mathbf{5 / 6}$ times per week Effect in percentual points With respect to the mean probability | $\begin{aligned} & -19.9053 \mathrm{pp} \\ & -71.1922 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & 16.1852 \mathrm{pp} \\ & 57.8869 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.6479 \mathrm{pp} \\ & -2.3173 \% \\ & \hline \end{aligned}$ | $\begin{gathered} 0.3922 \mathrm{pp} \\ 1.4028 \% \\ \hline \end{gathered}$ |
| Has not drunk during last 3 months or less than once per month <br> Effect in percentual points <br> With respect to the mean probability | $\begin{array}{r} 28.8476 \mathrm{pp} \\ 94.8309 \% \\ \hline \end{array}$ | $\begin{array}{r} -8.5222 \mathrm{pp} \\ -28.0151 \% \\ \hline \end{array}$ | $\begin{aligned} & 1.0238 \mathrm{pp} \\ & 3.3657 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.2453 \mathrm{pp} \\ & -0.8065 \% \\ & \hline \end{aligned}$ |
| Sedentary lifestyle Effect in percentual points With respect to the mean probability | $\begin{array}{r} -24.4977 \mathrm{pp} \\ -211.5522 \% \\ \hline \end{array}$ | $\begin{array}{r} 15.0736 \mathrm{pp} \\ 130.1692 \% \\ \hline \end{array}$ | $\begin{gathered} -0.7197 \mathrm{pp} \\ -6.2151 \% \\ \hline \end{gathered}$ | $\begin{aligned} & 1.0929 \mathrm{pp} \\ & 9.4381 \% \\ & \hline \end{aligned}$ |

Note: The effects for BSE are obtained after estimating equation (4) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B13 (and on Table D. 1 for first-stage regression). The effects for BME are obtained after estimating equation (8) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B13. Mean values for the dependent variables are reported on Table B6.

Table 3. Effect of one standard deviation increase (or decrease) of biased survival expectations (BSE) and biased meteorological expectations (BME) over health-related variables

|  | One standard deviation increase of BSE>0 | One standard deviation increase of BSE $<0$ | One standard deviation increase of BME $>0$ | One standard deviation increase of BME $<0$ |
| :---: | :---: | :---: | :---: | :---: |
| Normal weight Effect in percentual points With respect to the mean probability | $\begin{aligned} & 17.8893 \mathrm{pp} \\ & 47.5022 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -11.7843 \mathrm{pp} \\ & -31.2913 \% \\ & \hline \end{aligned}$ | $\begin{gathered} 0.0397 \mathrm{pp} \\ 0.1053 \% \\ \hline \end{gathered}$ | $\begin{array}{r} -0.0485 \mathrm{pp} \\ -0.1287 \% \\ \hline \end{array}$ |
| Obese or overweight <br> Effect in percentual points <br> With respect to the mean probability | $\begin{aligned} & -27.7656 p p \\ & -46.7670 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & 21.9060 \mathrm{pp} \\ & 36.8975 \% \\ & \hline \end{aligned}$ | $\begin{array}{r} -1.1183 \mathrm{pp} \\ -1.8836 \% \\ \hline \end{array}$ | $\begin{gathered} 0.8550 \mathrm{pp} \\ 1.4401 \% \\ \hline \end{gathered}$ |
| Ten word listing (first recall) <br> Effect in words <br> With respect to the mean probability | $\begin{array}{r} 0.5227 \text { words } \\ 10.4756 \% \\ \hline \end{array}$ | $\begin{array}{r} -0.1183 \text { words } \\ -2.3715 \% \\ \hline \end{array}$ | $\begin{array}{r} 0.0843 \text { words } \\ 1.6903 \% \\ \hline \end{array}$ | $\begin{array}{r} -0.0020 \text { words } \\ -0.0403 \% \\ \hline \end{array}$ |
| Ten word listing (second recall) <br> Effect in words <br> With respect to the mean probability | $\begin{array}{r} 0.5478 \text { words } \\ 15.3436 \% \\ \hline \end{array}$ | $\begin{array}{r} -0.3359 \text { words } \\ -9.4098 \% \\ \hline \end{array}$ | $\begin{array}{r} 0.0007 \text { words } \\ 0.0185 \% \\ \hline \end{array}$ | $\begin{array}{r} -0.0003 \text { words } \\ -0.0086 \% \\ \hline \end{array}$ |
| Maximum grip strength measure Effect in percentual points With respect to the mean probability | $\begin{array}{r} 1.4695 \text { kilos } \\ 4.3207 \% \\ \hline \end{array}$ | $\begin{array}{r} -1.0969 \mathrm{kilos} \\ -3.2253 \% \\ \hline \end{array}$ | $\begin{array}{r} 0.0147 \text { kilos } \\ 0.0432 \% \\ \hline \end{array}$ | $\begin{array}{r} -0.0060 \text { kilos } \\ -0.0178 \% \\ \hline \end{array}$ |

Note: BSE are obtained after estimating equation (4) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B14 (and on Table D. 1 for first-stage regression). BME are obtained after estimating equation (8) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B14. Mean values for the dependent variables are reported on Table B6.

Table 4. Effect of one standard deviation increase (or decrease) of biased survival expectations (BSE) and biased meteorological expectations (ME) over financial behaviour

|  | One standard deviation increase of BSE>0 | One standard deviation increase of BSE $<0$ | One standard deviation increase of BME $>0$ | One standard deviation increase of BME $<0$ |
| :---: | :---: | :---: | :---: | :---: |
| Bank accounts <br> Effect in percentual points <br> With respect to the mean probability | $\begin{aligned} & \text {-17.4381 pp } \\ & -18.6603 \% \\ & \hline \end{aligned}$ | $\begin{gathered} 16.4206 \mathrm{pp} \\ 17.5715 \% \\ \hline \end{gathered}$ | $\begin{array}{r} -0.1511 \mathrm{pp} \\ -0.1617 \% \\ \hline \end{array}$ | $\begin{aligned} & 0.0867 \mathrm{pp} \\ & 0.0927 \% \\ & \hline \end{aligned}$ |
| Government or corporate bonds Effect in percentual points With respect to the mean probability | $\begin{array}{r} 2.0500 \mathrm{pp} \\ 45.4539 \% \\ \hline \end{array}$ | $\begin{array}{r} -2.6957 \mathrm{pp} \\ -59.7725 \% \\ \hline \end{array}$ | $\begin{array}{r} 0.6083 \mathrm{pp} \\ 13.4869 \% \\ \hline \end{array}$ | $\begin{array}{r} -0.4730 \mathrm{pp} \\ -10.4882 \% \\ \hline \end{array}$ |
| Stocks or shares (listed or unlisted in stock market) Effect in percentual points With respect to the mean probability | $\begin{array}{r} -2.2562 \mathrm{pp} \\ -37.1694 \% \\ \hline \end{array}$ | $\begin{array}{r} 2.7745 \mathrm{pp} \\ 45.7084 \% \\ \hline \end{array}$ | $\begin{gathered} 0.0321 \mathrm{pp} \\ 0.5290 \% \\ \hline \end{gathered}$ | $\begin{aligned} & 0.1704 \mathrm{pp} \\ & 2.8073 \% \\ & \hline \end{aligned}$ |
| Mutual funds Effect in percentual points With respect to the mean probability | $\begin{array}{r} 3.7894 \mathrm{pp} \\ 50.1245 \% \\ \hline \end{array}$ | $\begin{gathered} -4.6898 \text { pp } \\ -62.0344 \% \\ \hline \end{gathered}$ | $\begin{gathered} 0.1266 \mathrm{pp} \\ 1.6741 \% \\ \hline \end{gathered}$ | $\begin{aligned} & -0.0323 \mathrm{pp} \\ & -0.4275 \% \\ & \hline \end{aligned}$ |
| Mutual funds: mostly in stocks Effect in percentual points With respect to the mean probability | $\begin{array}{r} -5.5313 \mathrm{pp} \\ -16.2208 \% \\ \hline \end{array}$ | $\begin{array}{r} 2.1874 \mathrm{pp} \\ 6.4146 \% \\ \hline \end{array}$ | $\begin{gathered} 0.1058 \mathrm{pp} \\ 0.3102 \% \\ \hline \end{gathered}$ | $\begin{gathered} 0.1263 \mathrm{pp} \\ 0.3705 \% \\ \hline \end{gathered}$ |
| Mutual funds: half in stocks, half in bonds <br> Effect in percentual points <br> With respect to the mean probability | $\begin{array}{r} -8.8961 \mathrm{pp} \\ -19.8044 \% \\ \hline \end{array}$ | $\begin{array}{r} 21.3028 \mathrm{pp} \\ 47.4238 \% \\ \hline \end{array}$ | $\begin{array}{r} -0.3325 \mathrm{pp} \\ -0.7401 \% \\ \hline \end{array}$ | $\begin{gathered} 0.2159 \mathrm{pp} \\ 0.4807 \% \\ \hline \end{gathered}$ |
| Mutual funds: mostly in bonds <br> Effect in percentual points <br> With respect to the mean probability | $\begin{gathered} 10.9752 \mathrm{pp} \\ 52.3128 \% \end{gathered}$ | $\begin{array}{r} -9.2257 \mathrm{pp} \\ -43.9736 \% \\ \hline \end{array}$ | $\begin{aligned} & 1.6812 \mathrm{pp} \\ & 8.0134 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.0676 \mathrm{pp} \\ & -0.3221 \% \\ & \hline \end{aligned}$ |
| Individual retirement accounts <br> Effect in percentual points <br> With respect to the mean probability | $\begin{array}{r} 5.0458 \mathrm{pp} \\ 69.2159 \% \\ \hline \end{array}$ | $\begin{array}{r} -5.8236 \mathrm{pp} \\ -79.8846 \% \\ \hline \end{array}$ | $\begin{gathered} 0.1965 \mathrm{pp} \\ 2.6949 \% \\ \hline \end{gathered}$ | $\begin{aligned} & -0.0426 \mathrm{pp} \\ & -0.5844 \% \\ & \hline \end{aligned}$ |
| Life insurance <br> Effect in percentual points <br> With respect to the mean probability | $\begin{array}{r} -5.2644 \mathrm{pp} \\ -43.6157 \% \\ \hline \end{array}$ | $\begin{array}{r} 6.1218 \mathrm{pp} \\ 50.7191 \% \\ \hline \end{array}$ | $\begin{array}{r} -2.7202 \mathrm{pp} \\ -22.5365 \% \\ \hline \end{array}$ | $\begin{array}{r} 1.7481 \mathrm{pp} \\ 14.4831 \% \\ \hline \end{array}$ |
| Mortgage <br> Effect in percentual points <br> With respect to the mean probability | $\begin{array}{r} 3.8137 \mathrm{pp} \\ 35.2791 \% \\ \hline \end{array}$ | $\begin{array}{r} -2.3360 \mathrm{pp} \\ -21.6092 \% \\ \hline \end{array}$ | $\begin{array}{r} 3.9915 \mathrm{pp} \\ 36.9237 \% \\ \hline \end{array}$ | $\begin{aligned} & -0.2909 \mathrm{pp} \\ & -2.6907 \% \\ & \hline \end{aligned}$ |
| Other debts (excluding mortgage) <br> Effect in percentual points <br> With respect to the mean probability | $\begin{array}{r} 3.4837 \mathrm{pp} \\ 30.5591 \% \end{array}$ | $\begin{array}{r} -2.8515 \mathrm{pp} \\ -25.0129 \% \\ \hline \end{array}$ | $\begin{gathered} 0.5289 \mathrm{pp} \\ 4.6396 \% \\ \hline \end{gathered}$ | $\begin{array}{r} -0.3158 \mathrm{pp} \\ -2.7705 \% \\ \hline \end{array}$ |
| Total wealth amount (1,000PPP) <br> Effect in 1,000 PPP <br> With respect to the mean probability | $\begin{aligned} & 5.2038 \mathrm{pp} \\ & 3.1103 \% \\ & \hline \end{aligned}$ | $\begin{array}{r} -11.5545 \mathrm{pp} \\ -6.9060 \% \\ \hline \end{array}$ | $\begin{gathered} 0.2418 \mathrm{pp} \\ 0.1445 \% \\ \hline \end{gathered}$ | $\begin{array}{r} -0.0220 \mathrm{pp} \\ -0.0132 \% \\ \hline \end{array}$ |

Note: BSE are obtained after estimating equation (4) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B15 (and on Table D. 1 for first-stage regression). The effects for BME are obtained after estimating equation (8) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B15. Mean values for the dependent variables are reported on Table B6.

Table 5. Effect of one standard deviation increase (or decrease) of biased survival expectations (BSE) and biased meteorological expectations (BME) over the amount of financial risk that the respondent is willing to take when making investments

|  | One standard deviation increase of BSE>0 | One standard deviation increase of BSE $<0$ | One standard deviation increase of BME $>0$ | One standard deviation increase of BME $<0$ |
| :---: | :---: | :---: | :---: | :---: |
| Take substantial risk expecting to earn substantial returns Effect in percentual points <br> With $t$ to the mean probability | $\begin{aligned} & 0.6016 \mathrm{pp} \\ & 6.9797 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.5048 \mathrm{pp} \\ & -5.8559 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & 0.1115 \mathrm{pp} \\ & 1.2929 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.2248 \mathrm{pp} \\ & -2.6074 \% \\ & \hline \end{aligned}$ |
| Take above average financial risks expecting to earn above average returns <br> Effect in percentual points <br> With $t$ to the mean probability | $\begin{aligned} & 25.8636 \text { pp } \\ & 83.1092 \% \end{aligned}$ | $\begin{aligned} & -17.1768 \mathrm{pp} \\ & -55.1955 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & 0.5233 \mathrm{pp} \\ & 1.6814 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.7786 \mathrm{pp} \\ & -2.5018 \% \\ & \hline \end{aligned}$ |
| Take average financial risks expecting to earn average returns <br> Effect in percentual points <br> With $t$ to the mean probability | $\begin{aligned} & -1.6327 \mathrm{pp} \\ & -6.5465 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & 1.8169 \mathrm{pp} \\ & 7.2849 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & -0.0756 \mathrm{pp} \\ & -0.3030 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & 0.1763 \mathrm{pp} \\ & 0.7068 \% \\ & \hline \end{aligned}$ |
| Not willing to take any financial risks Effect in percentual points With t to the mean probability | $\begin{aligned} & -13.5807 \mathrm{pp} \\ & -38.2341 \% \\ & \hline \end{aligned}$ | $\begin{gathered} 6.2775 \mathrm{pp} \\ 17.6732 \% \\ \hline \end{gathered}$ | $\begin{gathered} -0.2021 \mathrm{pp} \\ -0.5690 \% \\ \hline \end{gathered}$ | $\begin{aligned} & 0.2439 \mathrm{pp} \\ & 0.6865 \% \\ & \hline \end{aligned}$ |

Note: The effects for BSE are obtained after estimating equation (4) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B16 (and on Table D. 1 for first-stage regression). The effects for BME are obtained after estimating equation (8) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B16. Mean values for the dependent variables are reported on Table B6.

## Appendix for online publication only

## Appendix A. Definition of variables

Table A1 . Explanatory variables in the stock-sampling duration model

| Categories | Definition |
| :---: | :---: |
| Socio-demographic variables | o Age and age squared. <br> o Gender <br> o Level of education. Education is categorized in four levels using the 1997 International Standard Classification of Education (ISCED): primary, lower secondary, upper secondary, post-secondary non tertiary and tertiary education. <br> o Living alone <br> o Living in a rural area |
| Economic-related variables | o Household income refers to annual income (in euros) net of taxes collected by all household members. Income was adjusted for the purchasing power of different currencies using the PPP exchange rate of the year in which the "Main Questionnaire" was performed. Income was divided by the square root of the number of household members to adjust for household size. <br> o Household net worth is defined as the value of total assets (the total amount of financial assets plus the total amount of non-financial assets; note that this indicator only takes into account the value of dwellings from non-financial assets) minus the total value of outstanding liabilities as is expressed in the same terms as household income <br> o Being employed at the time of the survey. |
| Health status | o Self-reported health status ${ }^{51}$. We define five binary variables corresponding to the following categories: excellent, very good, good, fair and poor. <br> o Chronic conditions. We define seven binary variables corresponding to the following pathologies: Alzheimer, diabetes, cancer, high blood pressure, lung disease, stroke and heart attack. <br> o Number of days stayed at hospital during last year for surgery or medical tests. <br> o Number of days stayed at hospital during last year due to mental health problems. <br> o Number of days stayed in other institutions other than a hospital or a nursing home during last year (i.e., institutions for rehabilitation, convalescence) <br> o Number of visits to general practitioner during last year. |
| Smoking | [SHARE questions]: Have you ever smoked cigarettes, cigars or a pipe daily for a period of at least one year? and Do you smoke at the present time? <br> We define two binary variables: <br> o Current smoker: takes the value 1 if the respondent is smoking at the time of the survey, 0 otherwise. <br> o Past smoker: takes the value 1 if the respondent has smoked at least one year, but is not currently smoking, 0 otherwise. |
| Drinking alcohol | [SHARE questions]: I am now going to ask you a few questions about what you drink. During the last 3 months, how often have you drunk any alcoholic beverages, like beer, cider, wine, spirits or cocktails? We define two binary variables corresponding to the following categories <br> o Not having drunk during the last 3 months or less than one per month (teetotal) <br> o Drinking every day or 5-6 times per week |
| Sedentary lifestyle | [SHARE questions]: We would like to know about the type and amount of physical activity you do in your daily life. How often do you engage in vigorous physical activity (such as sports, heavy housework, or a job that involves physical labour) or in activities that require a moderate level of energy (such as gardening, cleaning the car, or doing a walk)? <br> We define a binary variable "sedentary lifestyle" that takes the value 1 if the respondent is engaged in a vigorous or moderate activity "one or three times a month" or "hardly ever, or never" •. |
| Weight | Using body mass index (BMI) provided by SHARE we define the following categories and include the variable "normal weight" in the regression. <br> o Underweight: if BMI is lower than 20. <br> o Normal weight: if BMI lies between 20 and 24.9 <br> o Overweight: if BMI lies between 25 and 29.9 <br> o Obese: if BMI is greater than 30 . |
| Memory test | Memory was tested using a standard version of two word-list learning tests (Brandt, 1991) with immediate and delayed recall. <br> The interviewer read a list of 10 words. In the immediate recall, participants are asked to recall as many words as possible in one minute, immediately after hearing them. In the delayed recall, participants are asked to recall as many words as possible in one minute, after several other interview questions. Each word correctly recalled scores 1 point. <br> Finally, the episodic memory score is calculated by adding up the number of target words recalled immediately, and the number of target words recalled after the delay. Thus, the score ranges between 0 and 20 with a high score indicating good cognitive function. |

[^25]Table A2. Explanatory variables in the first-stage regression for biased survival expectations

| Categories | Definition |
| :---: | :---: |
| Socio-demographic variables | o Age and age squared. <br> o Gender <br> o Level of education. Education is categorized in four levels using the 1997 International Standard Classification of Education (ISCED): primary, lower secondary, upper secondary, post-secondary non tertiary and tertiary education. <br> o Marital status (married/cohabiting, single, separated/divorced, widow) <br> o Living alone <br> o Living in a rural area |
| Economic-related variables | o Household income refers to annual income (in euros) net of taxes collected by all household members. Income was adjusted for the purchasing power of different currencies using the PPP exchange rate of the year in which the "Main Questionnaire" was performed. Income was divided by the square root of the number of household members to adjust for household size. <br> o Household net worth is defined as the value of total assets (the total amount of financial assets plus the total amount of non-financial assets; note that this indicator only takes into account the value of dwellings from non-financial assets) minus the total value of outstanding liabilities as is expressed in the same terms as household income <br> o Relation with economic activity at the time of the survey (employed, unemployed, retired, houseworking). |
| Health status | o Number of days stayed at hospital during last year for surgery or medical tests. <br> o Number of days stayed at hospital during last year due to mental health problems. <br> o Number of days stayed in other institutions other than a hospital or a nursing home during last year (i.e., institutions for rehabilitation, convalescence) <br> o Number of visits to general practitioner during last year. |

Table A3. Dependent variables in the biased survival models

| Categories | Definition |
| :---: | :---: |
| Smoking | As defined in Table A1 |
| Drinking alcohol | As defined in Table A1 |
| Sedentary lifestyle | As defined in Table A1 |
| Social contacts | [SHARE question] Which of the activities listed on this card - if any - have you done in the last twelve months? <br> We define a binary variable "social contacts" that takes the value 1 if the respondent has participated in the last month in voluntary or charity work, gone to sport club or social club, taken part in political/community/religious organization or attended to educational or training courses, 0 otherwise. |
| Locus of control | [SHARE question] How often do you feel that what happens to you is out of your control? <br> o Often <br> o Sometimes <br> o Rarely <br> o Never |
| Financial assets | [SHARE question]: The next questions ask about a number of different kinds of savings or investments that you may have. <br> We define five binary variables corresponding to the following assets: <br> o Bank accounts <br> o Government or corporate bonds <br> o Stocks or shares listed or unlisted in the stock market <br> o Funds or managed investment accounts: mostly in stocks, half in bonds and half in stocks, mostly in bonds. <br> o Retirement accounts <br> o Life insurance policies <br> [SHARE question] Do you have mortgages or loans on this property? <br> We define a binary variable "mortgage" that takes the value 1 if the respondent answers in the affirmative to previous question, 0 otherwise. <br> [SHARE question] The next question refers to money that you may owe, excluding mortgages or money owed on land, property or firms (if any). Looking at card 36, which of these types of debts do you currently have, if any? <br> We define the binary variable "Other debts (excluding mortgage)" if the respondent reports having any debt on cars and other vehicles (vans/motorcycles/boats, etc.), debt on credit cards / store cards, loans (from bank, building society or other financial institution), debts to relatives or friends, student loans or overdue bills (phone, electricity, heating, rent). |
| Risk taking attitude | [SHARE question]: When people invest their savings they can choose between assets that give low return with little risk to lose money, for instance a bank account or a safe bond, or assets with a high return but also a higher risk of losing, for instance stocks and shares. Which of the statements on the card comes closest to the amount of financial risk that you are willing to take when you save or make investments? <br> We define four binary variables corresponding to each answer: <br> o Take substantial financial risks expecting to earn substantial returns <br> o Take above average financial risks expecting to earn above average returns <br> o Take average financial risks expecting to earn average returns <br> o Not willing to take any financial risks. |

Table A4. Classification of countries
Countries have been group according to cultural similarities defined by the Lewis model (1996).

| Group | Countries | Categorization |
| :--- | :--- | :--- |
| Group 1 | Italy <br> Spain | Multi-active countries: warm, emotional, loquacious and impulsive |
| Group 2 | Belgium <br> France | A bit further from multi-active countries |
| Group 3 | Austria <br> Denmark | Share characteristics from multi-active countries and linear-active countries |
| Group 4 | Germany <br> Sweden <br> Switzerland | Linear-active countries, who are identified as factual, cool and decisive planners |

## Appendix B. Descriptive statistics

Figure B1. Survival expectations over time by wave (What are the chances that you will live to be age T or more?)


Note: This figure reports the probability of an individual subjective longevity attaining a target age. Using calibrated sampling weights. Target: $\mathrm{T}=75$ years if Age <65; T=80 years if $65 \leq$ Age $<70$; T=85 years if $70 \leq$ Age $<75$; T=90 years if $75 \leq$ Age $<80$; Individuals aged 80 or older have been collapsed at age 90 . In this case, we summarize the expectations of living to age 95 or older. Sample size: 21,384 for Wave 1; 21,741 for Wave 2; 32,405 for Wave 4; 42,495 for Wave 5; 38,295 for Wave 6.

Figure B2. Comparison of the subjective survival probability between follow-up respondents and non-follow-up respondents.


Each graph shows the kernel density function for the subjective survival probability reported by SHARE. Dashed line corresponds to the sample respondents who participate in two consecutive waves. Straight lines corresponds to the sample of respondents in one wave who disappear in the next one.

Figure B3. Kernel density function for parent's age of decease.


Note: Dashed line is used to represent the kernel density function of father's (mother's) age of decease for those fathers (mothers) who have died at the time of the survey. Straight line is used to represent the kernel density function of father's (mother's) age of decease for those fathers (mothers) who are still alive at time of the survey, and for whom age of decease has been predicted using multiple imputation. This technique allows predicting what age at decease for living parents would have been using information from the "Main Questionnaire". It requires two main assumptions: (i) the data must be missing at random, which is clearly the case because age at decease is missing for all living parents, and (ii) the reasons for the missing data must be captured by other variables that do not have missing values. As the age of decease is a continuous variable, and OLS imputation method has been chosen. We have used the deceased subsample and selected those variables for which parent's characteristics are also available. These variables are: (i) gender, (ii) age of the father/mother and age at the last survey for the deceased sample, (iii) number of children of the father/mother and number of children of the deceased, (iv) frequency of contact between respondent and father/mother and frequency of contact between the deceased and his/her children, (v) distance between father's/mother's household and respondent's household and distance between the deceased's household and his/her children, (vi) country and year fixed effects. To test the sensitivity of our results, we have selected five different random seed values, and added five different imputations to our main dataset. The results in these alternative cases were very similar to the original estimations.

Figures B4. Predicted Biased Survival Expectations by respondent's sex and age cohort and parents' living status.


Note : Predicted BSE are obtained from equation (3). The first two figures compare BSE according to respondent's age cohort and parents’
living status. The other four figures compare LO by gender for the same age cohort and parent's living status.

Table B1. Design of the samples.

|  | Wave 1 | Wave 2 | Wave 4 | Wave 5 | Wave 6 | Total |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Initial sample | 30,434 | 37,174 | 58,184 | 66,221 | 68,231 | 260,244 |
| Select countries interviewed in all waves ${ }^{\text {a }}$ | 22,119 | 22,381 | 33,418 | 43,491 | 38,979 | 160,388 |
| Select observations with not missing <br> calibrated sampling weights ${ }^{\text {b }}$ | 21,384 | 21,741 | 32,405 | 42,495 | 38,295 | 156,320 |

${ }^{\text {a }}$ Croatia not included in wave 1, 2, 4 and 5; Czech Republic not included in wave 1; Estonia not included in wave 1, 2 and 5; Greece not included in wave 4 and 5; Ireland not included in wave 1, 4, 5 and 6 ; Israel not included in wave 4 ; Hungary not included in wave 1, 2, 5 and 6; Luxembourg not included in wave 1, 2 and 4; Netherlands not included in wave 6; Poland not included in wave 1 and 5; Portugal not included in wave 1, 2 and 5; Slovenia not included in wave 1 and 2.
${ }^{\mathrm{b}}$ Calibrated sampling weights are missing for respondents younger than 50 years (i.e., age-ineligible partners of an age eligible respondent) and those with missing information on the set of calibration variables (i.e., age, gender and NUTS1 code).

|  | Wave 1 \& Wave 2 | Wave 2 \& Wave 3-4 | Wave 4 \& Wave 5 | Wave 5 \& Wave 6 | Total |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Merging consecutive waves of SHARE (obtained from last row of previous table) ${ }^{\text {c }}$ |  |  |  |  |  |
| Survivors sample | 14,493 | 13,933 | 24,515 | 32,559 | 85,500 |
| Deceased sample | 463 | 644 | 862 | 1,180 | 3,149 |
| Not follow-up respondents | 6,428 | 7,164 | 7,028 | 8,756 | 29,376 |
| Total | 21,384 | 21,741 | 32,405 | 42,495 | 118,025 |
| Report subjective survival probability |  |  |  |  |  |
| Survivors sample | 13,427 | 12,842 | 11,364 | 30,227 | 67,860 |
| Deceased sample | 372 | 524 | 284 | 860 | 2,040 |
| Not follow-up respondents | 5,733 | 6,346 | 3,738 | 7,877 | 23,694 |
| Total | 19,532 | 19,712 | 15,386 | 38,964 | 93,594 |
| Do not report subjective survival probability ${ }^{\text {d }}$ |  |  |  |  |  |
| Survivors sample | 1,066 | 1,091 | 13,151 | 2,332 | 17,640 |
| Deceased sample | 91 | 120 | 578 | 320 | 1,109 |
| Not follow-up respondents | 695 | 818 | 3,290 | 879 | 5,682 |
| Total | 1,852 | 2,029 | 17,019 | 3,531 | 24,431 |
| For the robustness checks |  |  |  |  |  |
| Excluding erroneous responses to probabilistic question |  |  |  |  |  |
| Survivors sample | 2,847 | 2,478 | 2,187 | 21,494 | 29,006 |
| Deceased sample | 169 | 178 | 106 | 724 | 1,177 |
| Total | 3,016 | 2,656 | 2,293 | 22,218 | 30,183 |
| Excluding focal responses for survival expectations' question: 0,50 or 100 |  |  |  |  |  |
| Survivors sample | 7,268 | 7,208 | 6,670 | 18,581 | 39,727 |
| Deceased sample | 198 | 253 | 163 | 552 | 1,166 |
| Total | 7,466 | 7,461 | 6,833 | 19,133 | 40,893 |

[^26]Table B2. Wave to wave retention rates of all wave 1 samples (\%)

|  | Wave 1\&2 | Wave 2\&3 | Wave 3\&4 | Wave 4\&5 | Wave 5\&6 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Austria | 74.4 | 71.3 | 74.3 | 78.3 | 81.4 |
| Beligium | 76.3 | 83.9 | 80.6 | 84.4 | 85.7 |
| Denmark | 77.0 | 80.2 | 85.2 | 89.6 | 88.3 |
| France | 67.0 | 76.2 | 82.4 | 72.5 | 71.1 |
| Germany | 55.1 | 73.6 | 77.6 | 68.3 | 89.5 |
| Italy | 71.4 | 87.1 | 84.8 | 88.0 | 89.3 |
| Spain | 68.6 | 83.3 | 80.1 | 89.2 | 88.2 |
| Sweden | 70.6 | 70.6 | 73.4 | 79.4 | 85.2 |
| Switzerland | 74.6 | 85.3 | 87.0 | 86.3 | 89.8 |
| All countries | 70.6 | 79.1 | 80.6 | 81.8 | 85.4 |

Source: Own work and Bergman et al. (2017)

Table B3. Test of equality of distributions. Comparison of the sample of follow-up respondents in two consecutive waves and the sample of non-follow-up respondents (participate in one wave, but disappear in the next one)

|  |  | Respondents in both waves | Non-respondents in the following wave | Test of equality of distributions ${ }^{\text {a }}$ |
| :---: | :---: | :---: | :---: | :---: |
| Waves 1 \&2 | Mean. <br> Std.Dev. <br> N | 62.44 (29.20) 12,890 | $\begin{gathered} \hline 57.89 \\ (30.35) \\ 6,642 \end{gathered}$ | $\begin{gathered} 1.257 \\ 0.2086 \end{gathered}$ |
| Waves 2 \& 4 | Mean. Std.Dev. N | $\begin{gathered} 46.02 \\ (32.58) \\ 993 \end{gathered}$ | 62.79 (29.68) 19,109 | $\begin{aligned} & -12.475 \\ & -12.475 \end{aligned}$ |
| Waves 4 \&5 | Mean. <br> Std.Dev. N |  | 60.61 (30.40) 4,421 | $\begin{gathered} 1.486 \\ 0.1374 \end{gathered}$ |
| Waves 5 \&6 | Mean. Std.Dev. N | (65.79) 28.59 $(29,298)$ | $\begin{gathered} 62.32 \\ (30.29) \\ 9.666 \end{gathered}$ | $\begin{gathered} 0.063 \\ 0.9500 \end{gathered}$ |
| All waves | Mean. <br> Std.Dev. $\mathrm{N}$ | $\begin{gathered} 64.54 \\ (28.96) \\ 53,242 \\ \hline \end{gathered}$ | $\begin{gathered} 61.61 \\ (30.07) \\ 35,838 \\ \hline \end{gathered}$ | $\begin{gathered} 1.212 \\ 0.2255 \end{gathered}$ |

[^27]Table B4. Survival expectations over time by wave (What are the chances that you will live to be age $\mathbf{T}$ or more?). By countries and groups of countries

|  | Target | T= 75 | T=80 | T=85 | T=90 | T= +90 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Austria | Mean | 73.94 | 65.67 | 58.28 | 46.99 | 36.18 |
|  | Std.Dev | (24.97) | (26.44) | (29.58) | (29.92) | (30.60) |
|  | N | 6,387 | 762 | 467 | 210 | 77 |
| Belgium | Mean | 67.50 | 62.68 | 55.63 | 44.70 | 33.86 |
|  | Std.Dev | (23.43) | (25.05) | (27.26) | (29.29) | (29.58) |
|  | N | 5,335 | 528 | 341 | 166 | 67 |
| Denmark | Mean | 78.86 | 74.27 | 66.85 | 53.87 | 37.82 |
|  | Std.Dev | (22.96) | (25.28) | (28.52) | (31.51) | (32.78) |
|  | N | 8,637 | 1,041 | 695 | 313 | 105 |
| France | Mean | 69.17 | 62.74 | 56.85 | 47.61 | 35.80 |
|  | Std.Dev | (25.53) | (27.25) | (27.32) | (29.52) | (29.76) |
|  | N | 6,635 | 833 | 566 | 284 | 112 |
| Germany | Mean | 69.54 | 60.83 | 54.32 | 41.47 | 30.93 |
|  | Std.Dev | (25.47) | (26.84) | (28.47) | (29.84) | (29.27) |
|  | N | 6,004 | 654 | 331 | 138 | 35 |
| Italy | Mean | 71.86 | 65.94 | 60.96 | 54.30 | 46.41 |
|  | Std.Dev | (26.64) | (27.83) | (29.88) | (31.82) | (32.97) |
|  | N | 6,274 | 826 | 443 | 156 | 54 |
| Spain | Mean | 72.75 | 65.19 | 59.57 | 48.97 | 38.45 |
|  | Std.Dev | (24.88) | (28.08) | (28.44) | (30.99) | (31.09) |
|  | N | 5,880 | 1,036 | 711 | 406 | 132 |
| Sweden | Mean |  | $67.50$ |  | $44.05$ |  |
|  | Std.Dev | (23.65) | $(26.53)$ | $(29.22)$ | $(31.22)$ | $(30.31)$ |
|  | N | 5,574 | 765 | 464 | 259 | 70 |
| Switzerland | Mean | 75.12 | 69.53 | 62.10 | 53.57 | 42.04 |
|  | Std.Dev | (22.25) | (24.26) | (26.25) | (29.22) | (30.98) |
|  | N | 4,883 | 620 | 386 | 178 | 60 |
| Group I |  |  |  |  |  |  |
|  | Std.Dev | $(26.03)$ | $(28.41)$ | $(29.24)$ | $(31.82)$ | $(32.30)$ |
|  | N | 8,085 | 2,184 | 2,169 | 1,783 | 1,697 |
| Group II | Mean | 67.57 | 61.68 | 55.62 | 46.26 | 33.97 |
|  | Std.Dev | (24.51) | (26.31) | (27.01) | (29.56) | (29.47) |
|  | N | 11,195 | 2,262 | 2,156 | 1,775 | 1,833 |
| Group III | Mean | 76.94 | 69.44 | $62.10$ | 50.90 | 38.12 |
|  | Std.Dev | (23.71) | (25.85) | (29.10) | (30.07) | (30.86) |
|  | N | 8,106 | 1,860 | 1,941 | 1,207 | 1,226 |
| Group IV |  | 72.99 | 66.46 | 58.23 | 47.11 | 34.85 |
|  | Std.Dev | (24.15) | (25.90) | (28.33) | (30.89) | (31.14) |
|  | N | 11,165 | 2,965 | 2,651 | 1,943 | 1,697 |
| All sample | Mean | 70.95 | 63.92 | 57.70 | 47.49 | 37.11 |
|  | Std.Dev | (25.47) | (27.27) | (28.69) | (30.81) | (31.22) |
|  | N | 55,609 | 7,065 | 4,404 | 2,110 | 712 |
| Test by country |  |  |  |  |  |  |
| Highest SSP <br> Test difference of means ${ }^{\text {a }}$ | Country F-test p-value | $\begin{gathered} \hline \text { Denmark } \\ \mathrm{F}_{1,68538}=530.98 \\ (0.0000) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Denmark } \\ \mathrm{F}_{1,17158}=175.80 \\ (0.0000) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Denmark } \\ \mathrm{F}_{1,16780}=105.35 \\ (0.0000) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Italy } \\ \mathrm{F}_{1,12851}=68.87 \\ (0.0000) \\ \hline \end{gathered}$ | Italy $\mathrm{F}_{1,13323}=101.86$ $(0.0000)$ |
| Lowest SSP <br> Test difference of means ${ }^{\text {b }}$ | Country <br> F-test <br> p-value | Belgium $\mathrm{F}_{1,68538}=124.53$ $(0.0000)$ | Germany $\mathrm{F}_{1,17158}=30.61$ $(0.0000)$ | Germany <br> $\mathrm{F}_{1,16780}=35.94$ <br> $(0.0000)$ | $\begin{gathered} \text { Germany } \\ \mathrm{F}_{1,12851}=73.36 \\ (0.0000) \\ \hline \end{gathered}$ | Sweden $\mathrm{F}_{1,13323}=91.54$ $(0.0000)$ |
| Test by group of countries |  |  |  |  |  |  |
| Highest SSP <br> Test difference of means ${ }^{\text {c }}$ | Group <br> F-test <br> p -value | $\begin{gathered} \text { Group III } \\ F_{1,3850}=116.89 \\ (0.0000) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Group III } \\ \mathrm{F}_{1,9270}=21.29 \\ (0.0000) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Group III } \\ \mathrm{F}_{1,8916}=6.37 \\ (0.0116) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Group I } \\ \mathrm{F}_{1,6707}=24.16 \\ (0.0000) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Group I } \\ \mathrm{F}_{1,6452}=48.89 \\ (0.0000) \\ \hline \end{gathered}$ |
| Lowest SSP <br> Test difference of means ${ }^{\text {d }}$ | Group <br> F-test <br> p-value | $\begin{gathered} \text { Group II } \\ \mathrm{F}_{1,38550}=39.66 \\ (0.0000) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Group II } \\ \mathrm{F}_{1,9270}=8.72 \\ (0.0039) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Group II } \\ \mathrm{F}_{1,8916}=3.03 \\ (0.0820) \\ \hline \end{gathered}$ | $\begin{gathered} \text { Group II } \\ \mathrm{F}_{1,6707}=0.01 \\ (0.8044) \\ \hline \end{gathered}$ | Group II $\mathrm{F}_{1,6452}=10.64$ $(0.0011)$ |

${ }^{\text {a }}$ Ho: mean(other countries)-mean(country with highest expectancy)=0
Ha: mean(other countries) < mean(country with highest expectancy)
${ }^{\text {b }}$ Ho: mean(other countries)-mean(country with lowest expectancy) $=0$
Ha: mean(other countries) > mean(country with lowest expectancy)
Ho: mean(other groups)-mean(group with the lowest expectancy) $=0$
Ha: mean(other groups) < mean(group with lowest expectancy)
Ho: mean(other groups)-mean(group with the lowest expectancy)=0
Ha: mean(other groups) > mean(group with lowest expectancy)
Note: Using calibrated sampling weights. Target: $T=75$ years if Age $<65$; $T=80$ years if $65 \leq$ Age $<70$; $T=85$ years if $70 \leq$ Age $<75$; $T=90$ years if $75 \leq$ Age
$<80$; Individuals aged 80 or older have been collapsed at age 90 . In this case, we summarize the expectations of living to age 95 or older.
Group I: Italy and Spain. Group II: Belgium and France. Group III: Austria and Denmark. Group IV: Germany, Sweden and Switzerland.

Table B5. Survival expectations by survival status and gender

|  | Target | T= 75 | T=80 | T=85 | T=90 | T= +90 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Survivors |  |  |  |  |  |  |
| Total | Mean | 70.79 | 64.07 | 58.40 | 48.56 | 38.47 |
|  | Std.Dev | (25.77) | (27.25) | (28.65) | (31.17) | (31.46) |
|  | N | 54,786 | 6,719 | 4,046 | 1,798 | 511 |
| Men | Mean | 70.58 | 64.32 | 58.79 | 50.01 | 40.79 |
|  | Std.Dev | (26.11) | (27.49) | (28.67) | (31.55) | (32.03) |
|  | N | 25,093 | 3,088 | 1,741 | 719 | 188 |
| Women | Mean | 70.98 | 63.84 | 58.07 | 47.52 | $37.13$ |
|  | Std.Dev | (25.44) | (27.03) | (28.64) | (30.85) | (31.06) |
|  | N | 29,693 | 3,631 | 2,305 | 1,079 | 323 |
| Test equality of means menwomen | F-test p-value | $\begin{gathered} \mathrm{F}_{1,29547}=0.65 \\ (0.4192) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,6940}=0.24 \\ (0.6276) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,6567}=0.45 \\ (0.5034) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,4678}=3.11 \\ (0.0779) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,3979}=5.29 \\ (0.0215) \\ \hline \end{gathered}$ |
| Deceased |  |  |  |  |  |  |
| Total | Mean | 57.23 | 47.09 | 44.01 | 44.11 | 29.80 |
|  | Std.Dev | (32.88) | (30.59) | (30.65) | (33.20) | (30.01) |
|  | N | 823 | 346 | 358 | 312 | 201 |
| Men | Mean | 56.06 | 46.53 | 47.49 | 47.34 | 28.69 |
|  | Std.Dev | (33.94) | (29.94) | (30.95) | (34.76) | (30.19) |
|  | N | 512 | 201 | 190 | 159 | 80 |
| Women | Mean | 59.90 | 47.95 | 39.39 | 40.23 | 30.60 |
|  | Std.Dev | (30.27) | (31.77) | (29.76) | (30.89) | (29.87) |
|  |  | $311$ | $145$ | $168$ | 153 | $121$ |
| Test equality of means menwomen | F-test <br> p-value | $\begin{gathered} \mathrm{F}_{1,372}=0.54 \\ (0.4622) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,194}=0.05 \\ (0.8272) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,291}=2.50 \\ (0.1148) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,360}=1.45 \\ (0.2296) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,118}=0.36 \\ (0.5484) \\ \hline \end{gathered}$ |
| Test equality of means survivors-deceased | F-test p-value | $\begin{gathered} \mathrm{F}_{129920}=27.34 \\ (0.0000) \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,7135}=28.39 \\ (0.0000) \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,6859}=30.34 \\ (0.0000) \\ \hline \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,5039}=1.92 \\ (0.1664) \end{gathered}$ | $\begin{gathered} \mathrm{F}_{1,4798}=22.84 \\ (0.0000) \\ \hline \end{gathered}$ |

Note: means and test of means have been computed using calibrated sampling weights.
Target: T=75 years if Age $<65$; $T=80$ years if $65 \leq$ Age $<70$; $T=85$ years if $70 \leq$ Age $<75$; T=90 years if $75 \leq$ Age $<80$; Individuals aged 80 or older have been collapsed at age 90 . In this case, we summarize the expectations of living to age 95 or older.

|  | Do not report subjective survival probabilities | Report subjective survival probabilities |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Not follow-up respondents | Final sample: Follow-up respondents |  |  |
|  |  |  | Total | Deceased sample | Survivor sample |
| N (all sample) | 24,431 | 23,694 | 69,900 | 2,040 | 67,860 |
| N (for the probability of sunny day) | - | - | 69,315 | 1,956 | 67,359 |
| N (respondents who participated in SHARELIFE) |  |  | 33,269 | 708 | 32,561 |
| Subjective survival probability (SSP) | - | 61.05 | 62.06 | 40.63 | 63.21 |
|  | - | (30.55) | (29.07) | (32.87) | (29.48) |
| Objective survival probability (OSP) | - | - | 44.07 | 17.18 | 47.99 |
|  | - |  | (28.97) | (24.54) | (29.35) |
| Biased survival expectations (SSP-OSP) | - | - | 19.99 | 23.44 | 15.21 |
|  | - | - | 36.00 | 40.43 | 35.86 |
| Biased survival expectations $>0$ ( $\mathrm{SSP}-\mathrm{OSP}>0$ ) | - | - | (24.26) | (27.44) | (24.15) |
|  | - | - | -20.42 | -12.02 | -20.66 |
| Biased survival expectations <0 (SSP- OSP $<0$ ) | - | - | (17.91) | (17.09) | (17.87) |
|  | - | - | (34.04) | (34.82) | (34.96) |
| Subjective probability of sunny day |  |  | 61.54 | 58.50 | 61.63 |
|  |  |  | (29.82) | (30.40) | (29.80) |
| Predicted meteorological probability of sunny day |  |  | 40.72 | 42.65 | 40.67 |
|  |  |  | (9.78) | (10.79) | (9.74) |
| Biased meteorological expectations |  |  | 20.82 | 15.85 | 20.96 |
|  |  |  | (29.78) | (30.67) | (29.74) |
| Biased meteorological expectations >0 |  |  | 33.80 | 31.57 | 33.86 |
|  |  |  | (18.90) | (18.30) | (18.91) |
| Biased meteorological expectations < 0 |  |  | -22.51 | -23.93 | -22.46 |
|  |  |  | (14.70) | (16.29) | (14.63) |
| Man | 43.36 | 46.48 | 46.22 | 52.91 | 45.99 |
| Woman | 56.64 | 53.52 | 53.78 | 47.09 | 54.01 |
| Age | 69.21 | 65.75 | 65.47 | 76.68 | 65.09 |
|  | (10.84) | (11.06) | (10.36) | (10.65) | (10.13) |
| Marital status ${ }^{\text {a }}$ |  |  |  |  |  |
| Single | 7.30 | 8.29 | 7.27 | 8.23 | 7.23 |
| Widow | 20.99 | 17.93 | 16.83 | 37.01 | 16.15 |
| Separated/divorced | 8.06 | 10.05 | 9.20 | 5.49 | 9.32 |
| Married/cohabiting | 63.65 | 63.73 | 66.70 | 49.27 | 67.30 |
| Education (ISCED 1997 levels) |  |  |  |  |  |
| None | 8.83 | 5.80 | 6.34 | 12.71 | 6.13 |
| Primary education | 24.88 | 20.78 | 24.52 | 40.30 | 23.99 |
| Lower secondary | 15.36 | 17.75 | 15.92 | 13.68 | 16.00 |
| Upper secondary | 31.21 | 35.57 | 31.51 | 22.70 | 31.80 |
| Post-secondary non tertiary | 2.06 | 2.06 | 2.40 | 1.51 | 2.43 |
| Tertiary education | 17.66 | 18.04 | 19.31 | 9.10 | 19.65 |
| Lives alone | 27.99 | 27.18 | 24.25 | 38.35 | 23.78 |
| Rural area, village, small town | 56.77 | 54.62 | 59.94 | 57.31 | 60.02 |
| Relation with economic activity |  |  |  |  |  |
| Employed | 20.86 | 29.88 | 28.59 | 4.72 | 29.39 |
| Unemployed | 2.68 | 3.41 | 4.05 | 1.40 | 4.14 |
| Retired | 54.12 | 49.20 | 50.34 | 70.40 | 49.66 |
| Disabled | 3.58 | 2.99 | 2.70 | 4.53 | 2.64 |
| Houseworking | 13.61 | 12.74 | 12.67 | 13.87 | 12.63 |
| Other | 5.15 | 1.78 | 1.65 | 5.08 | 1.54 |
| Adjusted wealth (1,000 PPP) ${ }^{\text {a }}$ |  |  |  |  |  |
|  | (254.10) | (336.06) | (286.63) | (191.16) | (289.19) |
| Adjusted income (1,000 PPP) ${ }^{\text {a }}$ | 18.00 | 22.62 | 19.99 | 15.43 | 20.15 |
|  | (21.25) | (74.04) | (25.60) | (21.45) | (25.72) |
| Body mass index |  |  |  |  |  |
| Underweight | 1.67 | 1.43 | 1.27 | 4.74 | 1.15 |
| Normal | 34.46 | 39.23 | 37.06 | 36.04 | 37.09 |
| Obese | 17.16 | 16.09 | 18.77 | 17.37 | 18.82 |
| Overweight | 39.04 | 40.81 | 40.60 | 35.74 | 40.77 |
| Missing | 7.67 | 2.44 | 2.30 | 6.11 | 2.17 |
| Sedentary (never moderate or vigorous physical activity) | 16.96 | 13.79 | 11.58 | 40.33 | 10.61 |
| Smoking |  |  |  |  |  |
| No, never | 37.79 | 51.20 | 48.82 | 46.62 | 48.89 |
| Not now | 48.56 | 25.50 | 24.17 | 22.79 | 27.14 |
| Yes | 13.65 | 23.30 | 27.01 | 30.59 | 23.97 |
| Drinking |  |  |  |  |  |
| Everyday or 5/6 times per week | 23.18 | 124.69 | 27.96 | 27.16 | 27.98 |
| Not drunk during last 3 months or les than once per month | 27.15 | 28.65 | 30.42 | 47.30 | 29.86 |
| Chronic illness |  |  |  |  |  |
| High blood pressure | 39.16 | 33.26 | 36.06 | 42.94 | 35.83 |
| Cancer | 5.35 | 5.47 | 5.11 | 14.14 | 4.81 |
| Stroke | 4.40 | 3.50 | 3.18 | 9.62 | 2.96 |
| Diabetes | 14.15 | 11.68 | 11.48 | 22.09 | 11.13 |
| Heart attack | 12.56 | 11.27 | 11.33 | 24.18 | 10.90 |
| Lung disease | 7.39 | 5.74 | 6.17 | 13.35 | 5.93 |
| Alzheimer | 3.43 | 1.06 | 0.68 | 4.19 | 0.56 |
| Days in hospital last year (surgery/medical tests) | 2.37 | 2.07 | 1.66 | 5.67 | 1.53 |
|  | (10.63) | (9.79) | (7.32) | (14.55) | (6.91) |
| Days in hospital last year (mental health) | (1.39 | 2.81 | 3.75 | 2.62 | 3.81 |
|  | (25.83) | (14.84) | (14.02) | (7.57) | (14.28) |
| Days in other institutions (last year) | $\begin{gathered} 26.10 \\ (28.60) \end{gathered}$ | $\begin{gathered} 46.33 \\ (81.87) \end{gathered}$ | $\begin{gathered} 27.44 \\ (31.63) \end{gathered}$ | $\begin{gathered} 54.16 \\ (67.43) \end{gathered}$ | $\begin{gathered} 25.42 \\ (26.06) \end{gathered}$ |


| Visits to medical practitioner last year | $\begin{gathered} 4.85 \\ (6.89) \end{gathered}$ | $\begin{gathered} 4.04 \\ (6.98) \end{gathered}$ | $\begin{gathered} 3.26 \\ (6.17) \end{gathered}$ | $\begin{gathered} 5.02 \\ (8.74) \end{gathered}$ | $\begin{gathered} 3.20 \\ (6.06) \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Blood test last year (participated in SHARELIFE) |  | . | 68.65 | 75.14 | 68.51 |
| Blood pressure checked last year (participated in SHARELIFE) |  | . | 69.12 | 74.15 | 69.01 |
| Depressed (last month) | 41.64 | 39.58 | 42.22 | 51.09 | 41.92 |
| Self-reported health status |  |  |  |  |  |
| Excellent | 4.51 | 6.96 | 7.53 | 2.37 | 7.70 |
| Very good | 12.10 | 16.12 | 16.52 | 5.88 | 16.88 |
| Good | 35.98 | 38.23 | 39.76 | 23.84 | 40.29 |
| Fair | 32.52 | 27.54 | 26.82 | 37.54 | 26.46 |
| Poor | 14.89 | 11.15 | 9.37 | 30.37 | 8.67 |
| Social contacts ${ }^{\text {b }}$ | 13.18 | 14.31 | 15.72 | 9.91 | 15.90 |
| Locus of control: Feels things are out of control |  |  |  |  |  |
| Often | 9.90 | 7.54 | 7.37 | 18.49 | 7.00 |
| Sometimes | 21.44 | 17.70 | 19.34 | 18.51 | 19.37 |
| Rarely | 22.22 | 19.05 | 21.44 | 16.96 | 21.60 |
| Never | 28.09 | 22.79 | 25.37 | 16.76 | 25.66 |
| Cognitive skills |  |  |  |  |  |
| Ten word listing (first recall) | 4.69 | 4.91 | 4.99 | 3.59 | 5.04 |
|  | (1.96) | (1.95) | (1.84) | (2.00) | (1.81) |
| Ten word listing (second recall) | 3.36 | 3.45 | 3.57 | 2.11 | 3.62 |
|  | (2.16) | (2.10) | (2.08) | (1.92) | (2.07) |
| Maximum grip strength | 32.34 | 34.34 | 34.01 | 27.67 | 34.19 |
|  | (12.16) | (12.72) | (12.31) | (11.85) | (12.27) |
| Financial assets |  |  |  |  |  |
| Bank accounts | 94.70 | 94.89 | 93.45 | 89.46 | 93.58 |
| Government or corporate bonds | 4.54 | 3.64 | 4.51 | 6.24 | 4.32 |
| Stocks or shares listed or unlisted in the stockmarket | 6.73 | 4.66 | 6.07 | 3.04 | 6.17 |
| Mutual funds or managed investment accounts | 6.84 | 4.37 | 7.56 | 2.98 | 6.36 |
| Mostly in stocks | 31.85 | 30.25 | 34.10 | 30.09 | 34.17 |
| Half in stocks and half in bonds | 41.87 | 40.09 | 44.92 | 39.82 | 45.01 |
| Mostly in bonds | 17.83 | 16.34 | 20.98 | 30.09 | 20.83 |
| Individual retirement accounts | 5.17 | 6.24 | 7.29 | 1.69 | 7.48 |
| Life insurance | 10.97 | 9.99 | 12.07 | 5.58 | 12.28 |
| Debts |  |  |  |  |  |
| Mortage | 7.54 | 8.86 | 10.81 | 3.31 | 11.06 |
| Other debts | 8.11 | 10.15 | 11.40 | 6.10 | 11.58 |
| Attitudes towards financial risk |  |  |  |  |  |
| Substantial risk expecting to earn substantial returns | 13.49 | 12.95 | 8.62 | 24.3 | 0.42 |
| Average fin. risk expecting to earn above average returns | 27.88 | 28.95 | 31.12 | 27.37 | 1.14 |
| Average financial risk expecting to earn average returns | 23.85 | 23.97 | 24.94 | 23.43 | 10.16 |
| Not willing to take any financial return | 34.78 | 34.13 | 35.32 | 24.90 | 35.33 |
| Countries |  |  |  |  |  |
| Austria | 0.96 | 3.16 | 3.37 | 3.17 | 3.38 |
| Germany | 34.56 | 34.81 | 25.21 | 21.29 | 25.34 |
| Sweden | 3.84 | 2.78 | 3.06 | 3.35 | 3.05 |
| Spain | 15.62 | 11.01 | 14.99 | 23.01 | 14.72 |
| Italy | 19.32 | 20.98 | 22.42 | 25.23 | 22.33 |
| France | 19.96 | 20.80 | 21.37 | 16.40 | 21.54 |
| Denmark | 1.96 | 1.25 | 2.00 | 2.17 | 1.99 |
| Switzerland | 1.27 | 2.00 | 3.34 | 2.18 | 3.37 |
| Belgium | 2.51 | 3.21 | 4.24 | 3.20 | 4.28 |

Income and wealth are expressed in 1,000 units of Purchase Power of Parity and adjusted by household sized (dividing wealth and income by the square root of the number of household members)
${ }^{\mathrm{b}}$ Respondent has participated in the last month in voluntary or charity work, gone to sport club or social club, taken part in political/community/religious organization or attended to educational or training courses.
Biased survival expectations: 50,058 observations for positive bias and 19,842 observations for negative bias
Biased meteorological expectations: 53,336 observations for positive bias and 15,979 observations for negative bias

Table B7. Descriptive statistics for reported and predicted parents' age of decease

|  | Father's age of decease |  |  | Mother's age of decease |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total | Men | Women | Total | Men | Women |
| Reported in SHARE |  |  |  |  |  |  |
| Mean | 78.58 | 78.73 | 78.46 | 81.97 | 81.93 | 82.01 |
| Std.dev | (8.78) | (8.77) | (8.78) | (8.65) | (8.58) | (8.71) |
| N | 49,423 | 22,801 | 26,622 | 42,823 | 19,768 | 23,055 |
| Predicted using multiple imputation |  |  |  |  |  |  |
| Mean | 82.74 | 82.79 | 82.69 | 84.59 | 84.74 | 84.47 |
| Std.dev | (5.19) | (4.98) | (5.36) | (6.21) | (6.19) | (6.23) |
| N | 20,477 | 9,170 | 11,307 | 27,077 | 12,203 | 14,874 |

Multiple imputation has been used to predict father's or mother's age of decease for those fathers or mother who were still alive at the time of the survey.

Table B8. Weibull estimations for the discrete-time hazard model

|  | Total sample |  | Excluding focal points |  | Excluding erroneous responses to probabilistic question |  | Including only individuals who have answered SHARELIFE |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | Hazard ratio | Coef. | Hazard ratio | Coef. | Hazard ratio | Coef. | Hazard ratio |
| Man | $\begin{aligned} & 0.633^{* * *} \\ & (0.083) \end{aligned}$ | $1.884^{* * *}$ | $\begin{aligned} & \hline 0.589^{* * *} \\ & (0.133) \end{aligned}$ | 1.802*** | $\begin{aligned} & 0.606^{* * *} \\ & (0.090) \end{aligned}$ | 1.833*** | $\begin{aligned} & 0.642^{* * *} \\ & (0.090) \end{aligned}$ | 1.895*** |
| Age | $\begin{aligned} & -0.049 \\ & (0.045) \end{aligned}$ | 0.952 | $\begin{aligned} & -0.012 \\ & (0.075) \end{aligned}$ | 0.988 | $\begin{aligned} & -0.045 \\ & (0.051) \end{aligned}$ | 0.956 | $\begin{aligned} & -0.057 \\ & (0.050) \end{aligned}$ | 0.956 |
| Age^2 | $\begin{aligned} & 0.001^{* * *} \\ & (0.000) \end{aligned}$ | $1.001^{* * *}$ | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | 1.001 | $\begin{aligned} & 0.001^{* *} \\ & (0.000) \end{aligned}$ | 1.001** | $\begin{aligned} & 0.003^{* * *} \\ & (0.001) \end{aligned}$ | 1.002*** |
| Employed | $\begin{aligned} & -0.653^{* * *} \\ & (0.183) \end{aligned}$ | 0.520*** | $\begin{aligned} & -0.449 * \\ & (0.238) \end{aligned}$ | 0.638* | $\begin{aligned} & -0.578^{* * *} \\ & (0.189) \end{aligned}$ | 0.561*** | $\begin{aligned} & -0.687^{* * *} \\ & (0.191) \end{aligned}$ | 0.529*** |
| Lives alone | $\begin{aligned} & 0.192^{* *} \\ & (0.087) \end{aligned}$ | 1.211** | $\begin{aligned} & 0.134 \\ & (0.149) \end{aligned}$ | 1.143 | $\begin{aligned} & 0.234^{* *} \\ & (0.098) \end{aligned}$ | 1.263** | $\begin{aligned} & 0.202^{* *} \\ & (0.094) \end{aligned}$ | 1.217** |
| Rural area, village, small town | $\begin{aligned} & -0.028 \\ & (0.071) \end{aligned}$ | 0.972 | $\begin{aligned} & -0.175 \\ & (0.116) \end{aligned}$ | 0.840 | $\begin{aligned} & -0.046 \\ & (0.078) \end{aligned}$ | 0.955 | $\begin{aligned} & -0.034 \\ & (0.076) \end{aligned}$ | 0.979 |
| Normal weight | $\begin{aligned} & -0.246 * * * \\ & (0.082) \end{aligned}$ | 0.782*** | $\begin{aligned} & -0.262^{* *} \\ & (0.32) \end{aligned}$ | 0.712** | $\begin{aligned} & -0.219^{* *} \\ & (0.090) \end{aligned}$ | 0.804** | $\begin{aligned} & -0.295^{* * *} \\ & (0.085) \end{aligned}$ | 0.801*** |
| High blood pressure | $\begin{aligned} & -0.087^{* *} \\ & (0.036) \end{aligned}$ | 0.916** | $\begin{aligned} & -0.225^{* *} \\ & (0.107) \end{aligned}$ | 0.799** | $\begin{aligned} & -0.109 * * \\ & (0.043) \end{aligned}$ | 0.896** | $\begin{aligned} & -0.091^{* *} \\ & (0.037) \end{aligned}$ | 0.918** |
| Heart attack | $\begin{aligned} & 0.253^{*} \\ & (0.127) \end{aligned}$ | 1.287* | $\begin{aligned} & 0.250^{*} \\ & (0.128) \end{aligned}$ | 1.284* | $\begin{aligned} & 0.241^{*} \\ & (0.119) \end{aligned}$ | 1.265* | $\begin{aligned} & 0.260^{*} \\ & (0.129) \end{aligned}$ | 1.291* |
| Cancer | $\begin{aligned} & 0.943^{* * *} \\ & (0.102) \end{aligned}$ | 2.566*** | $\begin{aligned} & 1.017^{* * *} \\ & (0.158) \end{aligned}$ | 2.765*** | $\begin{aligned} & 0.910^{* * *} \\ & (0.118) \end{aligned}$ | $2.484^{* * *}$ | $\begin{aligned} & 0.948^{* * *} \\ & (0.108) \end{aligned}$ | 2.568*** |
| Stroke | $\begin{aligned} & 0.434^{* * *} \\ & (0.122) \end{aligned}$ | 1.543*** | $\begin{aligned} & 0.300^{* *} \\ & (0.123) \end{aligned}$ | 1.350** | $\begin{aligned} & 0.575^{* * *} \\ & (0.143) \end{aligned}$ | 1.777*** | $\begin{aligned} & 0.440^{* * *} \\ & (0.126) \end{aligned}$ | 1.546*** |
| Diabetes | $\begin{aligned} & 0.346^{* * *} \\ & (0.089) \end{aligned}$ | 1.413*** | $\begin{aligned} & 0.533^{* * *} \\ & (0.156) \end{aligned}$ | 1.703*** | $\begin{aligned} & 0.375^{* * *} \\ & (0.101) \end{aligned}$ | 1.454*** | $\begin{aligned} & 0.340^{* * *} \\ & (0.088) \end{aligned}$ | 1.411*** |
| Lung disease | $\begin{aligned} & 0.248^{* *} \\ & (0.109) \end{aligned}$ | 1.281** | $\begin{aligned} & 0.364^{*} \\ & (0.193) \end{aligned}$ | 1.440* | $\begin{aligned} & 0.284^{* *} \\ & (0.121) \end{aligned}$ | 1.329** | $\begin{aligned} & 0.240^{* *} \\ & (0.101) \end{aligned}$ | 1.278** |
| Alzheimer | $\begin{aligned} & 0.461^{* * *} \\ & (0.157) \end{aligned}$ | 1.585*** | $\begin{aligned} & 0.210^{* * *} \\ & (0.051) \end{aligned}$ | 1.234*** | $\begin{aligned} & 0.411^{*} \\ & (0.216) \end{aligned}$ | 1.508* | $\begin{aligned} & 0.478^{* * *} \\ & (0.159) \end{aligned}$ | 1.592*** |
| Days at hospital (last year) | $\begin{aligned} & 0.012 * * * \\ & (0.001) \end{aligned}$ | 1.012*** | $\begin{aligned} & 0.015^{* * *} \\ & (0.002) \end{aligned}$ | 1.015*** | $\begin{aligned} & 0.014^{* * *} \\ & (0.002) \end{aligned}$ | 1.014*** | $\begin{aligned} & 0.014^{* * *} \\ & (0.001) \end{aligned}$ | 1.013*** |
| Days in hospital last year (mental health) | $\begin{aligned} & 0.0075 * * * \\ & (0.0015) \end{aligned}$ | 1.276*** | $\begin{aligned} & 0.0077 * * * \\ & (0.0016) \end{aligned}$ | 1.280*** | $\begin{aligned} & 0.0074^{* * *} \\ & (0.0014) \end{aligned}$ | 1.274*** | $\begin{aligned} & 0.0079^{* * *} \\ & (0.0016) \end{aligned}$ | 1.278*** |
| Days in other institutions (last year) | $\begin{aligned} & 0.0113^{* * *} \\ & (0.0021) \end{aligned}$ | 1.011*** | $\begin{aligned} & 0.0110^{* * *} \\ & (0.0019) \end{aligned}$ | 1.014*** | $\begin{aligned} & 0.0115^{* * *} \\ & (0.0020) \end{aligned}$ | 1.017*** | $\begin{aligned} & 0.0121^{* * *} \\ & (0.0023) \end{aligned}$ | 1.019*** |
| Visits to general practitioner (last year) | $\begin{aligned} & -0.002 \\ & (0.006) \end{aligned}$ | 0.998 | $\begin{aligned} & -0.002 \\ & (0.007) \end{aligned}$ | 0.998 | $\begin{aligned} & 0.002 \\ & (0.006) \end{aligned}$ | 1.002 | $\begin{aligned} & -0.003 \\ & (0.005) \end{aligned}$ | 0.999 |
| Sedentary lifestyle | $\begin{aligned} & 0.858^{* * *} \\ & (0.080) \end{aligned}$ | 2.357*** | $\begin{aligned} & 0.841^{* * *} \\ & (0.145) \end{aligned}$ | 2.318*** | $\begin{aligned} & 0.787 * * * \\ & (0.089) \end{aligned}$ | 2.196*** | $\begin{aligned} & 0.865^{* * *} \\ & (0.083) \end{aligned}$ | 2.360*** |
| Smokes now | $\begin{aligned} & 0.422^{* * *} \\ & (0.109) \end{aligned}$ | 1.524*** | $\begin{aligned} & 0.407^{* * *} \\ & (0.170) \end{aligned}$ | 1.487*** | $\begin{aligned} & 0.360^{* * *} \\ & (0.114) \end{aligned}$ | 1.434*** | $\begin{aligned} & 0.431^{* * *} \\ & (0.111) \end{aligned}$ | 1.527*** |
| Smoked before, not now | $\begin{aligned} & 0.313^{* * *} \\ & (0.108) \end{aligned}$ | 1.366*** | $\begin{aligned} & 0.327^{*} \\ & (0.145) \end{aligned}$ | 1.387** | $\begin{aligned} & 0.225^{*} \\ & (0.120) \end{aligned}$ | 1.253* | $\begin{aligned} & 0.315^{* * *} \\ & (0.111) \end{aligned}$ | 1.367*** |
| Drinks alcohol daily, 5-6 times/week | $\begin{aligned} & 0.215^{* *} \\ & (0.082) \end{aligned}$ | 1.234** | $\begin{aligned} & 0.227 \\ & (0.090) \end{aligned}$ | 1.246** | $\begin{aligned} & 0.240 \\ & (0.095) \end{aligned}$ | 1.267** | $\begin{aligned} & 0.217^{* *} \\ & (0.084) \end{aligned}$ | 1.235** |
| No education | $\begin{aligned} & 0.426^{* * *} \\ & (0.159) \end{aligned}$ | 1.531*** | $\begin{aligned} & 0.536^{* * *} \\ & (0.183) \end{aligned}$ | 1.710*** | $\begin{aligned} & 0.398^{* *} \\ & (0.168) \end{aligned}$ | 1.489** | $\begin{aligned} & 0.430^{* * *} \\ & (0.161) \end{aligned}$ | 1.533*** |
| Primary education | $\begin{aligned} & 0.454^{* * *} \\ & (0.132) \end{aligned}$ | 1.575*** | $\begin{aligned} & 0.390^{* *} \\ & (0.175) \end{aligned}$ | 1.477** | $\begin{aligned} & 0.427^{* * *} \\ & (0.135) \end{aligned}$ | 1.533*** | $\begin{aligned} & 0.459^{* * *} \\ & (0.134) \end{aligned}$ | 1.577*** |
| Lower secondary education | $\begin{aligned} & 0.187 \\ & (0.147) \end{aligned}$ | 1.205 | $\begin{aligned} & 0.256 \\ & (0.200) \end{aligned}$ | 1.292 | $\begin{aligned} & 0.182 \\ & (0.149) \end{aligned}$ | 1.199 | $\begin{aligned} & 0.195 \\ & (0.150) \end{aligned}$ | 1.208 |
| Upper secondary education | $\begin{aligned} & 0.300^{* *} \\ & (0.129) \end{aligned}$ | 1.349** | $\begin{aligned} & 0.214^{* *} \\ & (0.114) \end{aligned}$ | 1.239** | $\begin{aligned} & 0.335^{* * *} \\ & (0.127) \end{aligned}$ | 1.398*** | $\begin{aligned} & 0.308^{* *} \\ & (0.133) \end{aligned}$ | 1.351** |
| Post-secondary non tertiary | $\begin{aligned} & 0.387 \\ & (0.259) \end{aligned}$ | 1.472 | $\begin{aligned} & 0.401 \\ & (0.541) \end{aligned}$ | 1.579 | $\begin{aligned} & 0.302 \\ & (0.409) \end{aligned}$ | 1.381 | $\begin{aligned} & 0.393 \\ & (0.262) \end{aligned}$ | 1.476 |
| Adjusted wealth (1,000 PPP) | $\begin{aligned} & -0.411^{* *} \\ & (0.189) \end{aligned}$ | 0.663** | $\begin{aligned} & -0.561^{*} \\ & (0.227) \end{aligned}$ | 0.571** | $\begin{aligned} & -0.494^{* *} \\ & (0.213) \end{aligned}$ | 0.610** | $\begin{aligned} & -0.416^{* *} \\ & (0.193) \end{aligned}$ | 0.664** |
| Adjusted income (1,000 PPP) | $\begin{aligned} & -0.118 \\ & (1.346) \end{aligned}$ | 1.889 | $\begin{aligned} & -0.182 \\ & (1.152) \end{aligned}$ | 2.152 | $\begin{aligned} & 0.203 \\ & (1.331) \end{aligned}$ | $1.895$ | $\begin{aligned} & -0.121 \\ & (1.352) \end{aligned}$ | $1.890$ |
| Feeling depressed (last month) | $\begin{aligned} & 0.167 * * \\ & (0.072) \end{aligned}$ | 1.181** | $\begin{aligned} & 0.160^{* *} \\ & (0.078) \end{aligned}$ | 1.175** | $\begin{aligned} & 0.147^{* *} \\ & (0.050) \end{aligned}$ | 1.091** | $\begin{aligned} & 0.170^{* *} \\ & (0.071) \end{aligned}$ | $1.184^{* *}$ |
| Constant | $\begin{aligned} & -6.738^{* * *} \\ & (1.678) \end{aligned}$ | 0.001*** | $\begin{aligned} & -8.799^{* * *} \\ & (2.725) \\ & \hline \end{aligned}$ | 0.000*** | $\begin{aligned} & -6.912^{* * *} \\ & (1.883) \end{aligned}$ | 0.001*** | $\begin{aligned} & -6.777^{* * *} \\ & (1.685) \\ & \hline \end{aligned}$ | 0.002*** |
| N | 69.900 |  | 40,893 |  | 30,183 |  | 33,269 |  |
| Hazard rate <br> Wald chi2(38) p -value | $\begin{aligned} & 1.429 \\ & (0.057) \\ & 2385.24 \\ & (0.0000) \\ & \hline \end{aligned}$ |  | $\begin{gathered} 1.615 \\ (0.073) \\ 1876.38 \\ (0.0000) \\ \hline \end{gathered}$ |  | $\begin{gathered} 1.461 \\ (0.065) \\ 972.68 \\ (0.0000) \\ \hline \end{gathered}$ |  | $\begin{gathered} 1.477 \\ (0.060) \\ 1792.35 \\ (0.0000) \\ \hline \end{gathered}$ |  |

[^28]Table B9. Average daylight minutes per day ( $D M D_{m, r}$ )

|  | Jan | Feb | Mar | Apr | May | Jun | Jul | Aug | Sep | Oct | Nov | Dec |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Austria | 510 | 605 | 700 | 796 | 893 | 957 | 956 | 888 | 791 | 687 | 628 | 569 |
| Belgium | 488 | 593 | 698 | 803 | 910 | 982 | 980 | 904 | 797 | 683 | 618 | 553 |
| Denmark | 436 | 563 | 690 | 818 | 949 | 1,040 | 1,038 | 942 | 812 | 675 | 595 | 515 |
| France | 504 | 602 | 700 | 798 | 897 | 963 | 961 | 891 | 792 | 686 | 625 | 564 |
| Germany | 471 | 583 | 695 | 808 | 923 | 1,000 | 998 | 916 | 802 | 681 | 611 | 541 |
| Italy | 555 | 630 | 706 | 782 | 858 | 908 | 907 | 854 | 777 | 694 | 647 | 601 |
| Spain | 564 | 635 | 706 | 778 | 851 | 898 | 897 | 847 | 775 | 696 | 652 | 608 |
| Sweden | 382 | 532 | 682 | 833 | 986 | 1,101 | 1,098 | 978 | 825 | 668 | 572 | 477 |
| Switzerland | 520 | 611 | 702 | 793 | 885 | 946 | 944 | 880 | 787 | 688 | 632 | 576 |

Source: Gaisma

Table B10. Average sunlight minutes per day (SMD ${ }_{m, y, r}$ )

|  | Jan | Feb | Mar | Apr | May | Jun | Jul | Aug | Sep | Oct | Nov | Dec |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Austria | 142 | 197 | 268 | 322 | 385 | 405 | 442 | 412 | 347 | 280 | 161 | 127 |
| Belgium | 94 | 155 | 206 | 295 | 367 | 371 | 360 | 352 | 288 | 221 | 129 | 85 |
| Denmark | 73 | 133 | 215 | 350 | 463 | 453 | 443 | 412 | 284 | 183 | 112 | 66 |
| France | 172 | 219 | 309 | 384 | 437 | 503 | 551 | 492 | 406 | 294 | 203 | 165 |
| Germany | 100 | 156 | 229 | 319 | 403 | 419 | 423 | 402 | 313 | 225 | 121 | 90 |
| Italy | 221 | 248 | 310 | 374 | 452 | 516 | 584 | 535 | 432 | 336 | 243 | 204 |
| Spain | 269 | 296 | 363 | 407 | 493 | 560 | 618 | 571 | 453 | 366 | 293 | 252 |
| Sweden | 62 | 139 | 252 | 362 | 487 | 554 | 496 | 404 | 276 | 170 | 90 | 49 |
| Switzerland | 157 | 205 | 272 | 309 | 330 | 377 | 429 | 393 | 354 | 271 | 180 | 154 |

Source: World Meteorological Organization

Table B11. Average meteorological probability of sunny day (MetSD ${ }_{m, y, r}$ )

|  | Jan | Feb | Mar | Apr | May | Jun | Jul | Aug | Sep | Oct | Nov | Dec |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Austria | 0.28 | 32.63 | 38.35 | 40.45 | 43.16 | 42.27 | 46.28 | 46.39 | 43.81 | 40.77 | 25.61 | 22.40 |
| Belgium | 19.28 | 26.14 | 29.48 | 36.69 | 40.30 | 37.76 | 36.70 | 38.99 | 36.14 | 32.42 | 20.91 | 15.43 |
| Denmark | 16.75 | 23.66 | 31.23 | 42.75 | 48.80 | 43.54 | 42.72 | 43.76 | 34.98 | 27.07 | 18.78 | 12.75 |
| France | 34.13 | 36.40 | 44.10 | 48.10 | 48.77 | 52.27 | 57.32 | 55.17 | 51.26 | 42.90 | 32.46 | 29.21 |
| Germany | 21.29 | 26.83 | 33.00 | 39.45 | 43.61 | 41.94 | 42.42 | 43.84 | 39.05 | 33.00 | 19.74 | 16.62 |
| Italy | 39.81 | 39.29 | 43.87 | 47.84 | 52.69 | 56.83 | 64.34 | 62.59 | 55.55 | 48.36 | 37.56 | 33.91 |
| Spain | 47.78 | 46.60 | 51.43 | 52.26 | 57.95 | 62.41 | 68.87 | 67.40 | 58.43 | 52.60 | 44.99 | 41.50 |
| Sweden | 16.10 | 26.19 | 36.99 | 43.46 | 49.36 | 50.30 | 45.17 | 41.30 | 33.40 | 25.47 | 15.73 | 10.28 |
| Switzerland | 30.26 | 33.51 | 38.72 | 38.98 | 37.28 | 39.83 | 45.40 | 44.71 | 45.01 | 39.36 | 28.54 | 26.71 |

Source: own estimations using Tables B7 and B8.

Table B12. Comparison of biased expectation indicators

| By country |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Biased survival expectations OSP-SSP |  | Meteorological expectations |  | Comparison between indicators of expectations \% |  |  |  |
|  | Mean | Std.Dev. | Mean | Std.Dev. | LO $>0$ \&ME $>0$ | LO > 0 \& ME < 0 | LO < 0 \& ME > 0 | LO < 0 \& ME $<0$ |
| Total | 19.99 | 34.04 | 20.82 | 29.78 | 56.75 | 14.87 | 20.37 | 7.98 |
| Men | 27.62 | 33.93 | 21.18 | 29.76 | 61.74 | 16.37 | 15.46 | 6.37 |
| Women | 13.56 | 32.80 | 20.51 | 29.80 | 52.54 | 13.60 | 24.50 | 9.34 |
| Survivors | 19.81 | 34.03 | 20.96 | 29.74 | 56.84 | 14.75 | 20.42 | 7.98 |
| Deceased | 26.02 | 34.15 | 15.85 | 30.67 | 53.53 | 18.77 | 18.73 | 8.04 |
| Austria | 26.13 | 32.04 | 24.65 | 29.77 | 65.73 | 13.15 | 16.60 | 4.50 |
| Belgium | 10.96 | 32.43 | 22.91 | 28.66 | 49.03 | 12.87 | 27.57 | 10.52 |
| Denmark | 25.73 | 31.97 | 26.31 | 29.57 | 65.79 | 13.67 | 16.36 | 4.13 |
| France | 12.31 | 33.59 | 16.01 | 30.44 | 50.32 | 13.42 | 25.52 | 10.75 |
| Germany | 12.53 | 33.43 | 19.24 | 29.50 | 48.17 | 16.66 | 24.06 | 11.10 |
| Italy | 21.60 | 36.39 | 17.01 | 28.62 | 56.75 | 15.62 | 18.28 | 9.31 |
| Spain | 30.83 | 34.04 | 11.87 | 29.01 | 58.05 | 22.17 | 12.36 | 7.27 |
| Sweden | 23.63 | 33.13 | 26.38 | 29.94 | 61.55 | 13.57 | 18.28 | 6.52 |
| Switzerland | 20.55 | 32.69 | 24.91 | 29.09 | 60.76 | 12.58 | 21.09 | 5.53 |

Biased survival expectations: difference between objective survival probability and subjective survival probability
Biased meteorological expectations: difference between reported probability of sunny day and meteorological predicted probability of sunny day.

- By age cohort

| Age at baseline | N | LO>0 \& ME>00 |  |  | LO>0 \& ME<0 | LO<0 \& ME>0 |  | LO<0 \& ME<0 |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Mean LO | Mean ME | Mean LO | Mean ME | Mean LO | Mean ME | Mean LO | Mean ME |
| 50-64 | 55,609 | $\begin{gathered} 36,88 \\ (0,008) \end{gathered}$ | $\begin{gathered} 20,25 \\ (0,006) \end{gathered}$ | $\begin{gathered} 33,70 \\ (0,009) \end{gathered}$ | $\begin{gathered} \hline-22,78 \\ (0,008) \end{gathered}$ | $\begin{gathered} -21,01 \\ (0,013) \end{gathered}$ | $\begin{gathered} \hline 19,73 \\ (0,013) \end{gathered}$ | $\begin{aligned} & -21,00 \\ & (0,012) \end{aligned}$ | $\begin{gathered} \hline-27,24 \\ (0,016) \end{gathered}$ |
| 65-69 | 7,065 | $\begin{gathered} 34,68 \\ (0,006) \end{gathered}$ | $\begin{gathered} 40,14 \\ (0,007) \end{gathered}$ | $\begin{gathered} 31,14 \\ (0,013) \end{gathered}$ | $\begin{gathered} -19,53 \\ (0,011) \end{gathered}$ | $\begin{gathered} -20,77 \\ (0,010) \end{gathered}$ | $\begin{gathered} 36,50 \\ (0,015) \end{gathered}$ | $\begin{gathered} -22,25 \\ (0,016) \end{gathered}$ | $\begin{gathered} -21,86 \\ (0,018) \end{gathered}$ |
| 70-74 | 4,404 | $\begin{gathered} 34,41 \\ (0,007) \end{gathered}$ | $\begin{gathered} 44,77 \\ (0,009) \end{gathered}$ | $\begin{gathered} 31,07 \\ (0,016) \end{gathered}$ | $\begin{gathered} -20,23 \\ (0,014) \end{gathered}$ | $\begin{gathered} -20,95 \\ (0,011) \end{gathered}$ | $\begin{gathered} 40,43 \\ (0,017) \end{gathered}$ | $\begin{aligned} & -22,59 \\ & (0,020) \end{aligned}$ | $\begin{aligned} & -21,65 \\ & (0,023) \end{aligned}$ |
| 75-79 | 2,110 | $\begin{gathered} 32,55 \\ (0,008) \end{gathered}$ | $\begin{gathered} 47,04 \\ (0,011) \end{gathered}$ | $\begin{gathered} 31,97 \\ (0,018) \end{gathered}$ | $\begin{gathered} -15,48 \\ (0,014) \end{gathered}$ | $\begin{gathered} -21,34 \\ (0,013) \end{gathered}$ | $\begin{gathered} 41,15 \\ (0,021) \end{gathered}$ | $\begin{gathered} -24,40 \\ (0,021) \end{gathered}$ | $\begin{gathered} -16,35 \\ (0,021) \end{gathered}$ |
| 80-84 | 617 | $\begin{gathered} 31,55 \\ (0,010) \end{gathered}$ | $\begin{gathered} 45,46 \\ (0,015) \end{gathered}$ | $\begin{gathered} 31,59 \\ (0,021) \end{gathered}$ | $\begin{gathered} -10,45 \\ (0,015) \end{gathered}$ | $\begin{gathered} -22,44 \\ (0,015) \end{gathered}$ | $\begin{gathered} 39,15 \\ (0,028) \end{gathered}$ | $\begin{gathered} -23,06 \\ (0,027) \end{gathered}$ | $\begin{gathered} -10,49 \\ (0,025) \end{gathered}$ |
| 85-94 | 83 | $\begin{gathered} 29,74 \\ (0,015) \end{gathered}$ | $\begin{gathered} 42,81 \\ (0,023) \end{gathered}$ | $\begin{gathered} 30,74 \\ (0,031) \end{gathered}$ | $\begin{gathered} -3,55 \\ (0,011) \end{gathered}$ | $\begin{aligned} & -21,85 \\ & (0,024) \end{aligned}$ | $\begin{gathered} 38,26 \\ (0,044) \end{gathered}$ | $\begin{aligned} & -24,96 \\ & (0,353) \end{aligned}$ | $\begin{gathered} -5,47 \\ (0,077) \end{gathered}$ |
| 95-100 | 12 | $\begin{gathered} 30,75 \\ (0,025) \end{gathered}$ | $\begin{gathered} 45,86 \\ (0,037) \end{gathered}$ | $\begin{gathered} 30,13 \\ (0,079) \end{gathered}$ | $\begin{gathered} -0,42 \\ (0,005) \end{gathered}$ | $\begin{gathered} -24,37 \\ (0,045) \end{gathered}$ | $\begin{gathered} 36,65 \\ (0,066) \end{gathered}$ | $\begin{aligned} & -19,22 \\ & (0,272) \end{aligned}$ | $\begin{gathered} -2,08 \\ (0,029) \end{gathered}$ |
| All | 69,900 | $\begin{gathered} 30,13 \\ (0,008) \end{gathered}$ | $\begin{gathered} 34,52 \\ (0,011) \end{gathered}$ | $\begin{gathered} 28,56 \\ (0,016) \end{gathered}$ | $\begin{gathered} -13,32 \\ (0,011) \\ \hline \end{gathered}$ | $\begin{gathered} -19,59 \\ (0,013) \end{gathered}$ | $\begin{gathered} 30,89 \\ (0,021) \end{gathered}$ | $\begin{gathered} -19,94 \\ (0,059) \\ \hline \end{gathered}$ | $\begin{gathered} -15,29 \\ (0,027) \end{gathered}$ |

Table B13. Effect of biased survival expectations (LO) and biased meteorological expectation (ME) over lifestyle variables

|  | Using biased survival expectations <br> (IV- First-stage Table D.1) |  |  | Using meterological expectations (OLS) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | M1 | M2 | M3 | M1 | M2 | M3 |
| Smokes now |  |  |  |  |  |  |
| Biased expectations indicator > 0 | $\begin{aligned} & -0.6288^{* * *} \\ & (0.2291) \end{aligned}$ | $\begin{aligned} & \hline-0.6186^{* * *} \\ & (0.1981) \end{aligned}$ | $\begin{aligned} & \hline-0.5390^{* * *} \\ & (0.1924) \end{aligned}$ | $\begin{aligned} & -0.0117^{* *} \\ & (0.0046) \end{aligned}$ | $\begin{aligned} & -0.0111^{* *} \\ & (0.0046) \end{aligned}$ | $\begin{aligned} & -0.0102^{* *} \\ & (0.0046) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.4556 | 0.4565 | 0.4568 | 0.1546 | 0.1600 | 0.1634 |
| F | 994.820 | 987.895 | 990.151 | 1,338.748 | 972.895 | 901.434 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 42.65 | 41.38 | 49.68 |  |  |  |
| Durbin test chi2(1) /p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 24.69 | 25.88 | 26.02 |  |  |  |
| Sargan test chi2 | 2.56 | 2.12 | 1.96 |  |  |  |
| Smokes now |  |  |  |  |  |  |
| Biased expectations indicator < 0 | 0.8364*** | 0.8245*** | 0.8180*** | 0.0145*** | 0.0141** | 0.0135** |
|  | (0.2080) | (0.2007) | (0.2004) | (0.0048) | (0.0048) | (0.0048) |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2331 | 0.2337 | 0.2349 | 0.1801 | 0.1846 | 0.1854 |
| F | 353.321 | 329.516 | 328.189 | 5,090.501 | 3,446.851 | 3,032.646 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 41.90 | 42.05 | 42.12 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 25.13 | 25.41 | 25.67 |  |  |  |
| Sargan test chi2 | 2.77 | 2.87 | 2.80 |  |  |  |
| Has never smoked |  |  |  |  |  |  |
| Biased expectations indicator > 0 | 1.0967*** | $1.0443^{* *}$ | 1.0282** | 0.0068** | 0.0063** | 0.0062** |
|  | (0.4531) | (0.4532) | (0.4415) | (0.0027) | (0.0027) | (0.0027) |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2961 | 0.2968 | 0.2972 | 0.1003 | 0.1040 | 0.1061 |
| F | 250.651 | 273.376 | 246.077 | 2,583.187 | 1,767.512 | 1,581.170 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |


| Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $\begin{aligned} & 16.42 \\ & (0.0000) \\ & 24.56 \\ & 2.75 \\ & \hline \end{aligned}$ | $\begin{aligned} & 17.50 \\ & (0.0000) \\ & 25.22 \\ & 2.48 \\ & \hline \end{aligned}$ | $\begin{aligned} & 17.87 \\ & (0.0000) \\ & 5.98 \\ & 2.40 \\ & \hline \end{aligned}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Has never smoked |  |  |  |  |  |  |
| Biased expectations indicator $<0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $-1.8102^{* * *}$ $(0.4609)$ 19,842 0.1931 310.832 0.0000 15.14 $(0.0000)$ 23.26 2.32 | $-1.5526^{* * *}$ $(0.4682)$ 19,842 0.1937 377.704 0.0000 14.67 $(0.0000)$ 25.18 2.27 | $-1.4846 * * *$ $(0.4661)$ 19,842 0.1952 377.398 0.0000 16.01 $(0.0000)$ 26.48 2.16 | $-0.0232^{* *}$ $(0.0091)$ 15,979 0.1650 $4,581.661$ 0.0000 | $\begin{aligned} & \hline-0.0224^{* *} \\ & (0.0090) \\ & 15,979 \\ & 0.1842 \\ & 3,439.539 \\ & 0.0000 \end{aligned}$ | $\begin{aligned} & \hline-0.0222 * * \\ & (0.0097) \\ & 15,979 \\ & 0.2061 \\ & 1,581.170 \\ & 0.0000 \end{aligned}$ |
| Drinks everyday or 5/6 times per week |  |  |  |  |  |  |
| Biased expectations indicator $>0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $-0.9172^{* * *}$ $(0.2281)$ 50,058 0.2461 380.129 0.0000 22.63 $(0.0000)$ 28.49 1.65 | $-0.8690^{* * *}$ $(0.2104)$ 50,058 0.2470 476.718 0.0000 19.87 $(0.0000)$ 28.12 1.51 | $-0.8205^{* * *}$ $(0.2076)$ 50,058 0.2580 477.715 0.0000 20.08 $(0.0000)$ 30.05 1.33 | $-0.0353^{* * *}$ $(0.0065)$ 53,336 0.0980 384.154 0.0000 | $-0.0344^{* * *}$ $(0.0062)$ 53,336 0.1769 701.655 0.0000 | $-0.0343^{* * *}$ $(0.0062)$ 53,336 0.1790 690.038 0.0000 |
| Drinks everyday or 5/6 times per week |  |  |  |  |  |  |
| Biased expectations indicator $<0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $1.0217^{* * *}$ $(0.3048)$ 19,842 0.4412 438.442 0.0000 23.41 $(0.0000)$ 23.61 1.71 | $0.9397^{* * *}$ $(0.3012)$ 19,842 0.4424 442.966 0.0000 22.87 $(0.0000)$ 24.89 1.68 | $0.9042^{* * *}$ $(0.3093)$ 19,842 0.4437 443.865 0.0000 22.65 $(0.0000)$ 25.91 1.67 | $\begin{aligned} & \hline 0.0292^{* *} \\ & (0.0104) \\ & 15,979 \\ & 0.1705 \\ & 526.345 \\ & 0.0000 \end{aligned}$ | $\begin{aligned} & \hline 0.0272^{* *} \\ & (0.0104) \\ & 15,979 \\ & 0.1788 \\ & 389.475 \\ & 0.0000 \end{aligned}$ | $\begin{aligned} & \hline 0.0267 * * \\ & (0.0104) \\ & 15,979 \\ & 0.1814 \\ & 353.242 \\ & 0.0000 \end{aligned}$ |
| Has not drunk during last 3 months or less than once per month |  |  |  |  |  |  |
| Biased expectations indicator > 0 <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $1.2542^{* *}$ <br> $(0.5594)$ <br> 50,058 <br> 0.4981 <br> 166.635 <br> 0.0000 <br> 17.91 <br> $(0.0000)$ <br> 27.45 <br> 1.14 | $1.2264^{* *}$ <br> $(0.5553)$ <br> 50,058 <br> 0.5009 <br> 173.916 <br> 0.0000 <br> 15.43 <br> $(0.0000)$ <br> 28.50 <br> 1.05 | $1.1891^{* *}$ <br> $(0.5558)$ <br> 50,058 <br> 0.5129 <br> 171.262 <br> 0.0000 <br> 13.92 <br> $(0.0000)$ <br> 30.10 <br> 0.98 | $0.0554^{* *}$ $(0.0222)$ 53,336 0.2011 $1,746.392$ 0.0000 | 0.0545** $(0.0222)$ 53,336 0.2064 $1,184.407$ 0.0000 | $\begin{aligned} & \hline 0.0542 * * \\ & (0.0222) \\ & 53,336 \\ & 0.2077 \\ & 1,044.660 \\ & 0.0000 \end{aligned}$ |
| Has not drunk during last 3 months or less than once per month |  |  |  |  |  |  |
| Biased expectations indicator $<0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $-0.6053^{* * *}$ <br> $(0.1945)$ <br> 19,842 <br> 0.4740 <br> 106.225 <br> 0.0000 <br> 21.19 <br> $(0.0000)$ <br> 31.12 <br> 2.11 | $-0.4803^{* *}$ $(0.1822)$ 19,842 0.4953 102.295 0.0000 22.23 $(0.0000)$ 30.03 2.21 | $-0.4761^{* *}$ $(0.1435)$ 19,842 0.4979 100.993 0.0000 23.76 $(0.0000)$ 29.78 2.32 | $\begin{aligned} & \hline-0.0173^{* *} \\ & (0.0087) \\ & 15,979 \\ & 0.1995 \\ & 766.692 \\ & 0.0000 \end{aligned}$ | $-0.0172^{* *}$ $(0.0076)$ 15,979 0.1043 530.640 0.0000 | $-0.0167 * *$ $(0.0076)$ 15,979 0.1071 477.959 0.0000 |
| Sedentary lifestyle |  |  |  |  |  |  |
| Biased expectations indicator $>0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $-1.0182^{* *}$ $(0.4146)$ 50,058 0.2985 118.610 0.0000 13.26 $(0.0000)$ 24.58 1.06 | $-1.0135^{* *}$ $(0.4155)$ 50,058 0.3001 148.655 0.0000 13.95 $(0.0000)$ 25.87 0.87 | $-1.0098^{* *}$ $(0.4062)$ 50,058 0.3015 148.567 0.0000 17.57 $(0.0000)$ 27.49 0.75 | $-0.0495^{* *}$ $(0.0222)$ 53,336 0.1997 $1,730.474$ 0.0000 | $-0.0411^{*}$ $(0.0219)$ 53,336 0.2198 $1,283.449$ 0.0000 | $-0.0381^{*}$ $(0.0218)$ 53,336 0.2258 $1,162.097$ 0.0000 |
| Sedentary lifestyle |  |  |  |  |  |  |
| Biased expectations indicator $<0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $0.9650^{* *}$ $(0.3490)$ 19,842 0.1410 87.083 0.0000 17.79 $(0.0000)$ 25.67 1.26 |  <br> $0.8442^{* *}$ <br> $(0.3345)$ <br> 19,842 <br> 0.1435 <br> 195.183 <br> 0.0000 <br> 17.23 <br> $(0.0000)$ <br> 26.89 <br> 1.12 |  <br> $0.8421^{* *}$ <br> $(0.3015)$ <br> 19,842 <br> 0.1463 <br> 191.445 <br> 0.0000 <br> 17.48 <br> $(0.0000)$ <br> 27.05 <br> 1.19 | $0.0837 * * *$ $(0.0175)$ 15,979 0.1666 211.356 0.0000 | $0.0751^{* * *}$ $(0.0166)$ 15,979 0.2596 341.849 0.0000 | 0.0744*** $(0.0165)$ 15,979 0.2632 337.413 0.0000 |
| Education, marital status | Yes | Yes | Yes | Yes | Yes | Yes |


| Days at hospital for surgery/tests, due to mental problems, days at other institutions, visits to GP | No | Yes | Yes | No | Yes | Yes |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Relation with economic activity, lives alone, lives in rural area, income and wealth | No | No | Yes | No | No | Yes |
| Country fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |
| Time fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |

M1 includes optimistic indicator (after first-stage regression), gender, age, age squared, level of education (no education, primary education, lower secondary education, higher secondary education, post-secondary non-tertiary education and tertiary education (omitted)), marital status (married/cohabiting, separated/divorced, single and widow (omitted)), country fixed effects and wave fixed effects. M2 includes the same explanatory variables than M1, and additionally, number of days stayed at hospital during last year (surgery or medical tests), number of days stayed at hospital during last year (mental health), number of days stayed at other institutions during last year, number of visits to general practitioner during last year. M3 includes the same explanatory variables than M2, an additionally, relation with economic activity (employed, unemployed, retired, houseworking and other situations (omitted)), living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in 1,000 PPP). Bootstrap with 1,000 replications Robust standard errors between parenthesis. * statistically significant at $10 \%$; ** statistically significant at $5 \%$; *** statistically significant at $1 \%$.

Table B14. Effect of biased survival expectations (LO) and biased meteorological expectation (ME) over health-related variables

|  | Using biased survival expectations (IV- First-stage Table D.1) |  |  | Using meterological expectations (OLS) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | M1 | M2 | M3 | M1 | M2 | M3 |
| Normal weight |  |  |  |  |  |  |
| Biased expectations indicator > 0 | $\begin{aligned} & 0.9028^{* * *} \\ & (0.2279) \end{aligned}$ | $\begin{aligned} & 0.7616^{* * *} \\ & (0.2537) \end{aligned}$ | $\begin{aligned} & 0.7374^{* * *} \\ & (0.2315) \end{aligned}$ | $\begin{aligned} & \hline 0.0046 * * * \\ & (0.0011) \end{aligned}$ | $\begin{aligned} & 0.0033^{* * *} \\ & (0.0009) \end{aligned}$ | $\begin{aligned} & \hline 0.0021^{* *} \\ & (0.0007) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2860 | 0.2864 | 0.2878 | 0.1416 | 0.1831 | 0.1837 |
| F | 205.480 | 229.937 | 218.262 | 1,006.442 | 1,380.582 | 1,217.093 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 22.65 | 19.88 | 21.43 |  |  |  |
| Durbin test chi2(1) /p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 35.69 | 35.89 | 37.04 |  |  |  |
| Sargan test chi2 | 1.19 | 1.08 | 1.01 |  |  |  |
| Normal weight |  |  |  |  |  |  |
| Biased expectations indicator < 0 | $\begin{aligned} & -1.0093^{* * *} \\ & (0.2987) \end{aligned}$ | $\begin{aligned} & \hline-0.7323^{* *} \\ & (0.2736) \end{aligned}$ | $\begin{aligned} & \hline-0.6583^{* *} \\ & (0.2633) \end{aligned}$ | $\begin{aligned} & \hline-0.0057^{* *} \\ & (0.0019) \end{aligned}$ | $\begin{aligned} & -0.0040^{* *} \\ & (0.0017) \end{aligned}$ | $\begin{aligned} & \hline-0.0033^{* *} \\ & (0.0017) \end{aligned}$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2411 | 0.2416 | 0.2441 | 0.1218 | 0.1773 | 0.1778 |
| F | 152.947 | 218.386 | 207.788 | 516.924 | 1,275.193 | 1,124.469 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 27.86 | 26.81 | 26.50 |  |  |  |
| Durbin test chi2(1) /p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 33.51 | 32.90 | 32.12 |  |  |  |
| Sargan test chi2 | 1.32 | 1.28 | 1.11 |  |  |  |
| Obese or overweight |  |  |  |  |  |  |
| Biased expectations indicator > 0 | $\begin{aligned} & \hline-1.2110^{* *} \\ & (0.396) \end{aligned}$ | $\begin{aligned} & \hline-1.1909^{* *} \\ & (0.358) \end{aligned}$ | $\begin{aligned} & \hline-1.1445^{* *} \\ & (0.357) \end{aligned}$ | $\begin{aligned} & -0.0691^{* *} \\ & (0.0251) \end{aligned}$ | $-0.0604^{* *}$ (0.0248) | $\begin{aligned} & -0.0592^{* *} \\ & (0.0247) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.3527 | 0.4086 | 0.4286 | 0.2877 | 0.3022 | 0.3039 |
| F | 796.522 | 948.510 | 934.043 | 9,361.692 | 6,594.199 | 5,816.475 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 24.80 | 25.88 | 24.33 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 31.58 | 35.56 | 38.98 |  |  |  |
| Sargan test chi2 | 1.28 | 1.12 | 1.00 |  |  |  |
| Obese or overweight |  |  |  |  |  |  |
| Biased expectations indicator < 0 | $\begin{aligned} & 1.7996^{* * *} \\ & (0.554) \end{aligned}$ | $\begin{aligned} & 1.5021^{* * *} \\ & (0.559) \end{aligned}$ | $1.2238^{* *}$ (0.555) | $\begin{aligned} & \hline 0.0636 * * \\ & (0.0223) \end{aligned}$ | $0.0590^{* *}$ (0.0220) | $0.0582^{* *}$ $(0.0247)$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.3205 | 0.3221 | 0.3291 | 0.2685 | 0.2811 | 0.3039 |
| F | 595.790 | 694.857 | 688.146 | 8,506.316 | 5,955.134 | 5,816.475 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 29.56 | 28.75 | 28.21 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 30.41 | 28.61 | 28.19 |  |  |  |
| Sargan test chi2 | 1.42 | 1.37 | 1.31 |  |  |  |
| Ten word listing (first recall) |  |  |  |  |  |  |
| Biased expectations indicator > 0 | 3.3305*** | 2.8702*** | 2.1547*** | 0.7293*** | 0.4736*** | 0.4465*** |
|  | (0.711) | (0.703) | (0.686) | (0.1136) | (0.1102) | (0.1098) |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.6790 | 0.7183 | 0.7502 | 0.6596 | 0.6738 | 0.6754 |
| F | 2,583.271 | 3,144.535 | 3,236.081 | 6,518.774 | 6,417.712 | 6,267.699 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 40.87 | 47.56 | 45.31 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 23.14 | 25.68 | 27.48 |  |  |  |
| Sargan test chi2 | 1.43 | 1.26 | 1.17 |  |  |  |
| Ten word listing (first recall) |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | $\begin{aligned} & \hline-0.6965^{* * *} \\ & (0.1586) \end{aligned}$ | $\begin{aligned} & \hline-0.6801^{* * *} \\ & (0.1949) \end{aligned}$ | $\begin{aligned} & \hline-0.6611^{* * *} \\ & (0.1853) \end{aligned}$ | $\begin{aligned} & \hline-0.0156^{* *} \\ & (0.0042) \end{aligned}$ | $\begin{aligned} & \hline-0.0139^{* * *} \\ & (0.0042) \end{aligned}$ | $\begin{aligned} & \hline-0.0137 * * * \\ & (0.0042) \end{aligned}$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.4407 | 0.4558 | 0.4667 | 0.3424 | 0.3453 | 0.3455 |
| F | 966.729 | 966.972 | 899.419 | 1,840.370 | 1,721.758 | 1,670.683 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 43.31 | 41.12 | 40.98 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 27.89 | 26.89 | 26.40 |  |  |  |
| Sargan test chi2 | 1.89 | 1.88 | 1.80 |  |  |  |
| Ten Word listing (second recall) |  |  |  |  |  |  |
| Biased expectations indicator >0 | $2.7465^{* * *}$ |  |  |  |  |  |
|  | (0.7131) | (0.7081) | (0.7064) | (0.0006) | (0.0005) | (0.0005) |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.4214 | 0.4503 | 0.4871 | 0.3404 | 0.3878 | 0.3885 |
| F | 32.152 | 54.933 | 53.322 | 292.132 | 438.188 | 387.022 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 55.21 | 53.24 | 57.88 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 24.89 | 25.67 | 27.89 |  |  |  |
| Sargan test chi2 | 1.62 | 1.53 | 1.35 |  |  |  |
| Ten Word listing (second recall) |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | $\begin{aligned} & \hline-2.4732 * * * \\ & (0.7882) \end{aligned}$ | $\begin{aligned} & \hline-1.8469 * * \\ & (0.7348) \end{aligned}$ | $\begin{aligned} & \hline-1.8767^{* *} \\ & (0.6536) \end{aligned}$ | $\begin{aligned} & \hline-0.0034^{* *} \\ & (0.0012) \end{aligned}$ | $-0.0027^{* *}$ (0.0010) | $\begin{aligned} & \hline-0.0021^{* *} \\ & (0.0009) \end{aligned}$ |


| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $\mathrm{R}^{2}$ | 0.4811 | 0.4821 | 0.4832 | 0.3203 | 0.3822 | 0.3834 |
| F | 17.590 | 46.451 | 44.790 | 143.582 | 407.714 | 362.376 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 57.27 | 56.34 | 56.15 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 27.90 | 25.91 | 26.39 |  |  |  |
| Sargan test chi2 | 1.89 | 1.84 | 1.75 |  |  |  |
| Maximum grip strength measure |  |  |  |  |  |  |
| Biased expectations indicator $>0$ | 7.0176*** | 6.0616** | 6.0572** | 0.1120*** | 0.0828** | 0.0778** |
|  | (2.093) | (2.647) | (2.168) | (0.0352) | (0.0339) | (0.0338) |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.5446 | 0.5446 | 0.5456 | 0.3205 | 0.3427 | 0.3446 |
| F | 226.097 | 243.278 | 236.869 | 3,272.160 | 2,375.145 | 2,095.338 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 47.12 | 46.49 | 45.43 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 27.04 | 2890 | 30.27 |  |  |  |
| Sargan test chi2 | 1.07 | 0.98 | 0.81 |  |  |  |
| Maximum grip strength measure |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | -7.8529** | -7.250** | -6.1280** | -0.0643*** | -0.0551** | -0.0411** |
|  | (2.626) | (2.535) | (2.214) | (0.0189) | (0.0179) | (0.0178) |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.4549 | 0.5121 | 0.5171 | 0.2947 | 0.3110 | 0.3135 |
| F | 261.062 | 266.289 | 260.737 | 2,899.058 | 2,056.523 | 1,819.981 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 43.14 | 42.25 | 41.36 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 27.56 | 26.57 | 26.13 |  |  |  |
| Sargan test chi2 | 1.15 | 1.09 | 1.03 |  |  |  |
| Education, marital status | Yes | Yes | Yes | Yes | Yes | Yes |
| Days at hospital for surgery/tests, due to mental problems, days at other institutions, visits to GP | No | Yes | Yes | No | Yes | Yes |
| Relation with economic activity, lives alone, lives in rural area, income and wealth | No | No | Yes | No | No | Yes |
| Country fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |
| Time fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |

M1 includes optimistic indicator (after first-stage regression), gender, age, age squared, level of education (no education, primary education, lower secondary education, higher secondary education, post-secondary non-tertiary education and tertiary education (omitted)), marital status (married/cohabiting, separated/divorced, single and widow (omitted)), country fixed effects and wave fixed effects. M2 includes the same explanatory variables than M1, and additionally, number of days stayed at hospital during last year (surgery or medical tests), number of days stayed at hospital during last year (mental health), number of days stayed at other institutions during last year, number of visits to general practitioner during last year. M3 includes the same explanatory variables than M2, an additionally, relation with economic activity (employed, unemployed, retired, houseworking and other situations (omitted)), living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in 1,000 PPP).

Table B15. Effect of biased survival expectations (LO) and biased meteorological expectation (ME) over investment behaviour

|  | Using biased survival expectations <br> (IV- First-stage Table D.1) |  |  | Using meterological expectations (OLS) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | M1 | M2 | M3 | M1 | M2 | M3 |
| Bank accounts |  |  |  |  |  |  |
| ```Biased expectations indicator >0 N \(\mathrm{R}^{2}\) F p Durbin test chi2(1)/p-value Stock \& Yogo's test (critical value at 5\%: 19.93) Sargan test chi2``` | $-0.7648^{* * *}$ <br> $(0.1925)$ <br> 50.058 <br> 0.1415 <br> 472.601 <br> 0.0000 <br> 20.87 <br> $(0.0000)$ <br> 25.05 <br> 1.98 | $-0.7531^{* * *}$ $(0.2208)$ 50,058 0.1463 458.540 0.0000 20.78 $(0.0000)$ 24.93 1.90 | $-0.7188^{* * *}$ <br> $(0.2139)$ <br> 50.058 <br> 0.1555 <br> 439.838 <br> 0.0000 <br> 20.93 <br> $(0.0000)$ <br> 25.56 <br> 2.01 | $-0.0101^{* * *}$ $(0.0034)$ 53,336 0.0933 $2,385.473$ 0.0000 | $-0.0078^{* *}$ $(0.0034)$ 53,336 0.0960 $1,617.503$ 0.0000 | $-0.0080^{* *}$ $(0.0033)$ 53,336 0.1000 $1,481.197$ 0.0000 |
| Bank accounts |  |  |  |  |  |  |
| Biased expectations indicator < 0 <br> N <br> R ${ }^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $0.9519 * * *$ $(0.1790)$ 19,842 0.1526 106.557 0.0000 39.90 $(0.0000)$ 31.57 1.38 | 0.9277*** $(0.2059)$ 19,842 0.1544. 108.158 0.0000 39.13 $(0.0000)$ 33.21 1.26 |  <br> $0.9175^{* * *}$ <br> $(0.2001)$ <br> $19,842$. <br> 01605 <br> 100.499 <br> 0.0000 <br> 40.15 <br> $(0.0000)$ <br> 34.07 <br> 1.11 | $\begin{aligned} & \hline 0.0079 * * * \\ & (0.0021) \\ & 15,979 \\ & 0.0421 \\ & 1,017.662 \\ & 0.0000 \end{aligned}$ | $\begin{aligned} & 0.0058^{* *} \\ & (0.0021) \\ & 15,979 \\ & 0.0439 \\ & 698.768 \\ & 0.0000 \end{aligned}$ | $\begin{aligned} & \hline 0.0059 * * \\ & (0.0021) \\ & 15,979 \\ & 0.0443 \\ & 618.141 \\ & 0.0000 \end{aligned}$ |
| Government or corporate bonds |  |  |  |  |  |  |
| ```Biased expectations indicator >0 N R2 F p Durbin test chi2(1)/p-value Stock & Yogo's test (critical value at 5%: 19.93) Sargan test chi2``` | $0.0950^{* *}$ $(0.0378)$ 50,058 0.1890 308.975 0.0000 19.68 $(0.0000)$ 27.48 1.85 | $0.0942^{* *}$ $(0.0392)$ 50,058 0.2089 306.252 0.0000 21.15 $(0.0000)$ 28.49 1.74 | $0.0845 * *$ $(0.0325)$ 50,058 0.2080 289.932 0.0000 22.43 $(0.0000)$ 29.59 1.63 | 0.0363*** $(0.0066)$ 53,336 0.1012 $2,608.859$ 0.0000 | $\begin{aligned} & \hline 0.0355^{* * *} \\ & (0.0066) \\ & 53,336 \\ & 0.1043 \\ & 1,772.503 \\ & 0.0000 \end{aligned}$ | $\begin{aligned} & \hline 0.0352^{* * *} \\ & (0.0066) \\ & 53,336 \\ & 0.1046 \\ & 1,557.134 \\ & 0.0000 \end{aligned}$ |
| Government or corporate bonds |  |  |  |  |  |  |
| ```Biased expectations indicator < 0 N R2 F p Durbin test chi2(1)/p-value Stock & Yogo's test (critical value at 5%: 19.93) Sargan test chi2``` | -0.1690** $(0.0833)$ 19,842 0.1844 198.929 0.0000 41.06 $(0.0000)$ 30.75 1.42 | -0.1538** $(0.0776)$ 19,842 0,1890 192.150 0.0000 41.12 $(0.0000)$ 33.07 1.46 | $-0.1506 * *$ $(0.0773)$ 19,842 0.1904 185.140 0.0000 41.55 $(0.0000)$ 30.96 1.53 | $-0.0443^{* * *}$ $(0.0107)$ 15,979 0.0774 $1,944.628$ 0.0000 | $\begin{aligned} & \hline 0-.04010^{* * *} \\ & (0.0107) \\ & 15,979 \\ & 0.0807 \\ & 1,337.700 \\ & 0.0000 \end{aligned}$ | $\begin{aligned} & \hline-0.0322^{* * *} \\ & (0.0066) \\ & 15,979 \\ & 0.1046 \\ & 1,557.134 \\ & 0.0000 \end{aligned}$ |
| Stocks or shares (listed or unlisted in stockmarket) |  |  |  |  |  |  |
| ```Biased expectations indicator >0 N \(\mathrm{R}^{2}\) F p Durbin test chi2(1)/p-value Stock \& Yogo's test (critical value at 5\%: 19.93) Sargan test chi2``` | -0.0987** $(0.0041)$ 50,058 0.2451 52.101 0.0000 14.89 $(0.0000)$ 27.89 1.33 | $-0.0951^{* *}$ $(0.0037)$ 50,058 0.2472 44.870 0.0000 16.56 $(0.0000)$ 28.12 1.21 | $-0.0930^{* *}$ $(0.0034)$ 50,058 0.2497. 44.399 0.0000 17.25 $(0.0000)$ 30.15 1.17 | $-0.0021^{* *}$ $(0.0007)$ 53,336 0.1721 79.600 0.0000 | $-0.0018^{* *}$ $(0.0008)$ 53,336 0.1737 73.689 0.0000 | $\begin{aligned} & \hline 0.0017^{* *} \\ & (0.0008) \\ & 53,336 \\ & 0.1747 \\ & 71.668 \\ & 0.0000 \end{aligned}$ |
| Stocks or shares (listed or unlisted in stockmarket) |  |  |  |  |  |  |
| ```Biased expectations indicator < 0 N \(\mathrm{R}^{2}\) F p Durbin test chi2(1)/p-value Stock \& Yogo's test (critical value at 5\%: 19.93) Sargan test chi2``` | $0.1620 * * *$ $(0.0527)$ 19.842 0.2003 127.306 0.0000 15.78 $(0.0000)$ 34.13 0.71 | $\begin{aligned} & \hline 0.1584^{* * *} \\ & (0.0510) \\ & 19,842 \\ & 0.8033 \\ & 124.511 \\ & 0.0000 \\ & 17.42 \\ & (0.0000) \\ & 35.63 \\ & 0.68 \\ & \hline \end{aligned}$ | $0.1550 * * *$ $(0.0492)$ 19,842 0.2075 125.508 0.0000 18.16 $(0.0000)$ 35.99 0.57 | 0.0178*** $(0.0050)$ 15,979 0.0941 39.764 0.0000 | $\begin{aligned} & \hline 0.0121^{* *} \\ & (0.0051) \\ & 15,979 \\ & 0.0978 \\ & 38.007 \\ & 0.0000 \end{aligned}$ | $\begin{aligned} & \hline 0.0116 * * \\ & (0.00351) \\ & 15,979 \\ & 0.0989 \\ & 37.154 \\ & 0.0000 \end{aligned}$ |
| Mutual funds or managed investment accounts |  |  |  |  |  |  |
| ```Biased expectations indicator >0 N \(\mathrm{R}^{2}\) F p Durbin test chi2(1)/p-value Stock \& Yogo's test (critical value at 5\%: 19.93) Sargan test chi2``` | 0.1746*** $(0.0504)$ 50,058 0.1835 129.095 0.0000 17.56 $(0.0000)$ 26.87 1.25 | 0.1675*** $(0.0490)$ 50,058 0.1905 128.607 0.0000 19.85 $(0.0000)$ 27.12 1.16 | $0.1562^{* * *}$ $(0.0487)$ 50,058 0.1990 128.803 0.0000 17.02 $(0.0000)$ 27.89 1.08 | 0.0086*** $(0.0021)$ 53,336 0.1518 126.050 0.0000 | 0.0072*** $(0.0022)$ 53,336 0.1566 90.690 0.0000 | $0.0067 * * *$ $(0.0022)$ 53,336 0.1583 81.767 0.0000 |
| Mutual funds or managed investment accounts |  |  |  |  |  |  |


| Biased expectations indicator < 0 | $\begin{aligned} & \hline-0.2877 * * * * \\ & (0.0884) \end{aligned}$ | $\begin{aligned} & -0.2812 * * * \\ & (0.0862) \end{aligned}$ | $\begin{aligned} & \hline-0.2620^{* * *} \\ & (0.0830) \end{aligned}$ | $\begin{aligned} & \hline-0.0029 * * * \\ & (0.0009) \end{aligned}$ | $\begin{aligned} & \hline-0.0028^{* * *} \\ & (0.0009) \end{aligned}$ | $\begin{aligned} & \hline-0.0022^{* *} \\ & (0.0009) \end{aligned}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2187 | 0.2261 | 0.2280 | 0.1292 | 0.1318 | 0.1439 |
| F | 268.113 | 261.238 | 283.621 | 3,439.609 | 2,312.261 | 2,240.060 |
| p | 0.0000 | 0.0000 |  | 0.0000 | 0.0000 | 0.0000 |
|  | 19.12 | 21.10 | 23.95 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 35.52 | 36.41 | 37.14 |  |  |  |
| Sargan test chi2 | 0.93 | 0.95 | 1.07 |  |  |  |
| Mutual funds: mostly in stocks |  |  |  |  |  |  |
| Biased expectations indicator > 0 | $-0.2659 * *$ $(0.1266)$ | $-0.2438 * *$ <br> (0.1081) | $\begin{aligned} & -0.2280^{* *} \\ & (0.1007) \end{aligned}$ | $\begin{aligned} & \hline 0.0077^{* *} \\ & (0.0028) \end{aligned}$ | $\begin{aligned} & \hline 0.0060^{* *} \\ & (0.0028) \end{aligned}$ | 0.0056* |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2145 | 0.2054 | 0.2015 | 0.1099 | 0.1106 | 0.1142 |
| F | 321.422 | 316.970 | 290.943 | 2,862.599 | 1,894.342 | 1,717.918 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 23.05 | 23.12 | 23.23 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 20.45 | 21.07 | 20.88 |  |  |  |
| Sargan test chi2 | 0.89 | 0.83 | 0.85 |  |  |  |
| Mutual funds: mostly in stocks |  |  |  |  |  |  |
| Biased expectations indicator < 0 | $\begin{aligned} & \hline 0.2115 * * * \\ & (0.0574) \end{aligned}$ | $\begin{aligned} & \hline 0.1898 * * * \\ & (0.0583) \end{aligned}$ | $\begin{aligned} & 0.1222 * * \\ & (0.0581) \end{aligned}$ | $\begin{aligned} & 0.0110^{* *} \\ & (0.0039) \end{aligned}$ | $\begin{aligned} & 0.0086^{* *} \\ & (0.0039) \end{aligned}$ | $\begin{aligned} & 0.0086 * * \\ & (0.0038) \end{aligned}$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2587 | 0.2632 | 0.12730 | 0.1587 | 0.1600 | 0.1142 |
| F | 499.326 | 492.068 | 466.396 | 4,372.017 | 2,901.209 | 1,717.918 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 29.88 | 32.14 | 32.45 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 27.15 | 27.78 | 28.99 |  |  |  |
| Sargan test chi2 | 0.76 | 0.62 | 0.53 |  |  |  |
| Mutual funds: half in stocks, half in bonds |  |  |  |  |  |  |
| Biased expectations indicator > 0 | $\begin{aligned} & -0.3843^{* * *} \\ & (0.0992) \end{aligned}$ | $\begin{aligned} & -0.3430^{* * *} \\ & (0.0953) \end{aligned}$ | $\begin{aligned} & -0.3667 * * * \\ & (0.0886) \end{aligned}$ | $\begin{aligned} & -0.0192^{* * *} \\ & (0.0075) \end{aligned}$ | $\begin{aligned} & \hline-0.0173^{* *} \\ & (0.0075) \end{aligned}$ | $\begin{aligned} & \hline-0.0176^{* *} \\ & (0.0075) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2233 | 0.2422 | 0.2350 | 0.0915 | 0.0940 | 0.0952 |
| F | 243.503 | 245.547 | 231.024 | 355.720 | 338.892 | 332.992 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 18.45 | 18.70 | 18.23 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 26.58 | 26.90 | 26.05 |  |  |  |
| Sargan test chi2 | 1.22 | 1.18 | 1.27 |  |  |  |
| Mutual funds: half in stocks, half in bonds |  |  |  |  |  |  |
| Biased expectations indicator < 0 | $\begin{aligned} & 1.2878 * * * \\ & (0.3270) \end{aligned}$ | $\begin{aligned} & 1.2716 * * * \\ & (0.3706) \end{aligned}$ | $\begin{aligned} & 1.1901^{* *} \\ & (0.3493) \end{aligned}$ | $\begin{aligned} & \hline 0.0190^{* * *} \\ & (0.0064) \end{aligned}$ | $\begin{aligned} & \hline 0.0149^{* *} \\ & (0.0064) \end{aligned}$ | $\begin{aligned} & \hline 0.0147 * * \\ & (0.0064) \end{aligned}$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2487 | 0.2501 | 0.2516 | 0.1207 | 0.1254 | 0.1275 |
| F | 152.517 | 170.215 | 159.798 | 484.941 | 468.263 | 462.255 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 47.34 | 44.06 | 42.39 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 30.12 | 29.51 | 30.62 |  |  |  |
| Sargan test chi2 | 1.90 | 1.82 | 1.93 |  |  |  |
| Mutual funds: mostly in bonds |  |  |  |  |  |  |
| Biased expectations indicator > 0 | $0.5507 * *$ | $0.4990^{* *}$ | $0.4524^{* *}$ | $\begin{aligned} & \hline 0.0908^{* * *} \\ & (0.0137) \end{aligned}$ | $0.0889 * * *$ |  |
|  | (0.2064) 50,058 | (0.2185) 50,058 | (0.1908) $50,058$ |  | (0.0137) 53,336 | $(0.0136)$ |
| $\mathrm{R}^{2}$ | 0.1843 | 0.2053 | 0.2244 | 0.1193939.418 | 0.1220 | 53,336 0.1295 |
| F | 161.446 | 153.819 | 156.363 |  | 633.141 | $\begin{aligned} & 0.1295 \\ & 592.782 \end{aligned}$ |
| p | 0.0000 | 0.0000 | 0.000028.41 | 0.0000 | 0.0000 | 0.0000 |
|  | 29.25 | 29.32 |  |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 23.25 | 22.72 | 23.75 |  |  |  |
| Sargan test chi2 | 1.87 | 1.80 | 1.90 |  |  |  |
| Mutual funds: mostly in bonds |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | $\begin{aligned} & \hline-0.5231^{* *} \\ & (0.2320) \end{aligned}$ | $\begin{aligned} & \hline-0.5166 * * \\ & (0.2246) \end{aligned}$ | $\begin{aligned} & \hline-0.5154^{* *} \\ & (0.2109) \end{aligned}$ | $\begin{aligned} & -0.0059^{* *} \\ & (0.0014) \end{aligned}$ | $\begin{aligned} & -0.0039 * * \\ & (0.0014) \end{aligned}$ |  |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.1723 | 0.1789 | 0.1801 | 0.0458 | 0.0491 | 0.0507 |
| F | 143.468 | 141.115 | 140.335 | 333.1690.0000 | 235.3760.0000 | $\begin{aligned} & 212.725 \\ & 0.0000 \end{aligned}$ |
| p | 0.0000 | 0.0000 | 0.0000 |  |  |  |
|  | 32.27 | 28.86 | 29.84 | 0.0000 | 0.0000 | $0.0000$ |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 32.62 | 33.20 | 32.09 |  |  |  |
| Sargan test chi2 | 1.23 | 1.19 | 1.28 |  |  |  |
| Individual retirement accounts |  |  |  |  |  |  |
| Biased expectations indicator > 0 | 0.2314*** | 0.2219*** | 0.2080*** | 0.0106** | 0.0107** | 0.0104** |
|  | (0.0422) | (0.0475) | (0.0481) | (0.0046) | (0.0046) | (0.0046) |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2541 | 0.2574 | 0.2614 | 0.1067 | 0.1097 | 0.1106 |
| F | 177.505 | 175.603 | 172.946 | 828.622 | 561.533 | 495.501 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 24.67 | 22.05 | 22.97 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |


| Stock \& Yogo’s test (critical value at 5\%: 19.93) Sargan test chi2 | $\begin{aligned} & 21.58 \\ & 0.74 \\ & \hline \end{aligned}$ | $\begin{aligned} & 22.18 \\ & 0.61 \\ & \hline \end{aligned}$ | $\begin{aligned} & 22.94 \\ & 0.52 \\ & \hline \end{aligned}$ |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Individual retirement accounts |  |  |  |  |  |  |
| Biased expectations indicator < 0 | $\begin{aligned} & -0.3414^{* * *} \\ & (0.0989) \end{aligned}$ | $\begin{aligned} & -0.371^{* * *} \\ & (0.0950) \end{aligned}$ | $\begin{aligned} & -0.3253^{* * *} \\ & (0.0917) \end{aligned}$ | $\begin{aligned} & \hline-0.0034^{* *} \\ & (0.0015) \end{aligned}$ | $\begin{aligned} & \hline-0.0033^{* *} \\ & (0.0015) \end{aligned}$ | $\begin{aligned} & \hline-0.0029^{* *} \\ & (0.0014) \end{aligned}$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.1745 | 0.1781 | 0.1856 | 0.0946 | 0.0987 | 0.0995 |
| F | 144.797 | 146.473 | 144.632 | 724.619 | 498.811 | 440.278 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 36.35 | 36.46 | 35.36 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 26.56 | 27.42 | 27.35 |  |  |  |
| Sargan test chi2 | 0.90 | 0.84 | 0.86 |  |  |  |
| Life insurance |  |  |  |  |  |  |
| Biased expectations indicator > 0 | $\begin{aligned} & -0.2772^{* * *} \\ & (0.0726) \end{aligned}$ | $\begin{aligned} & \hline-0.2369^{* *} \\ & (0.0711) \end{aligned}$ | $\begin{aligned} & -0.2170^{* * *} \\ & (0.0647) \end{aligned}$ | $\begin{aligned} & \hline-0.1599 * * * \\ & (0.0321) \end{aligned}$ | $\begin{aligned} & -0.1481^{* * *} \\ & (0.0330) \end{aligned}$ | $\begin{aligned} & \hline-0.1440 * * * \\ & (0.0329) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.3414 | 0.4183 | 0.4537 | 0.2046 | 0.2100 | 0.2152 |
| F | 39.322 | 37.313 | 37.724 | 31.454 | 29.283 | 28.989 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 37.58 | 35.76 | 34.14 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 28.65 | 29.21 | 29.60 |  |  |  |
| Sargan test chi2 | 0.73 | 0.85 | 0.97 |  |  |  |
| Life insurance |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | 0.3679** | 0.3570** | 0.3420** | 0.1214** | 0.1176** | 0.1190** |
|  | (0.1265) | (0.1210) | (0.1217) | (0.0396) | (0.0305) | (0.0305) |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2561 | 0.2502 | 0.2531 | 0.1401 | 0.1472 | 0.1527 |
| F | 92.556 | 95.944 | 98.928 | 19.925 | 19.019 | 19.056 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 27.31 | 27.95 | 27.39 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 35.20 | 35.59 | 37.23 |  |  |  |
| Sargan test chi2 | 1.27 | 1.18 | 1.09 |  |  |  |
| Mortgage |  |  |  |  |  |  |
| Biased expectations indicator $>0$ | 0.1637*** | 0.1614*** | 0.1572*** | 0.2039** | 0.2104** | 0.2113** |
|  | (0.0463) | (0.0458) | (0.0453) | (0.0835) | (0.0842) | (0.0843) |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2012 | 0.2078 | 0.2112 | 0.1580 | 0.1637 | 0.1660 |
| F | 116.258 | 114.058 | 112.582 | 42.050 | 30.301 | 27.479 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 21.41 | 23.35 | 23.87 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 25.89 | 26.90 | 27.08 |  |  |  |
| Sargan test chi2 | 0.71 | 0.68 | 0.57 |  |  |  |
| Mortgage |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | -0.1504** | -0.1444** | -0.1305* | $-0.0201 * * *$ | -0.0208*** | -0.0198*** |
|  | (0.0629) | (0.0599) | (0.0567) | (0.0056) | (0.0056) | (0.0055) |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2341 | 0.2374 | 0.2463 | 0.1250 | 0.1291 | 0.1401 |
| F | 103.852 | 102.026 | 106.996 | 991.108 | 675.336 | 649.295 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 20.31 | 23.31 | 20.11 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 36.41 | 37.02 | 40.06 |  |  |  |
| Sargan test chi2 | 1.36 | 1.23 | 1.19 |  |  |  |
| Other debts (excluding mortgage) |  |  |  |  |  |  |
| Biased expectations indicator > 0 | $\begin{aligned} & \hline 0.1598^{* *} \\ & (0.0584) \end{aligned}$ | $\begin{aligned} & \hline 0.1563^{* *} \\ & (0.0520) \end{aligned}$ | $\begin{aligned} & \hline 0.1436^{* *} \\ & (0.0527) \end{aligned}$ | $\begin{aligned} & \hline 0.0313^{* * *} \\ & (0.0081) \end{aligned}$ | $\begin{aligned} & \hline 0.0290^{* * *} \\ & (0.0081) \end{aligned}$ | $\begin{aligned} & \hline 0.0280^{* * *} \\ & (0.0081) \end{aligned}$ |
| N | (0.0584) 50,058 | (0.0520) | (0.0527) | 53,336 | (0.0081) | (0.0081) |
| $\mathrm{R}^{2}$ | 0.2150 | 0.2192 | 0.2280 | 0.1145 | 0.1162 | 0.1193 |
| F | 131.755 | 123.257 | 123.566 | 896.992 | 598.711 | 539.706 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 32.56 | 31.41 | 32.09 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 24.52 | 25.72 | 26.19 |  |  |  |
| Sargan test chi2 | 1.37 | 1.25 | 1.10 |  |  |  |
| Other debts (excluding mortgage) |  |  |  |  |  |  |
| Biased expectations indicator < 0 | $\begin{aligned} & \hline-0.1640^{* *} \\ & (0.0769) \end{aligned}$ | $\begin{aligned} & \hline-0.1623^{* *} \\ & (0.0712) \end{aligned}$ | $\begin{aligned} & \hline-0.1593^{* *} \\ & (0.0725) \end{aligned}$ | $\begin{aligned} & \hline-0.0209^{* *} \\ & (0.0076) \end{aligned}$ | $\begin{aligned} & \hline-0.0198^{* *} \\ & (0.0076) \end{aligned}$ | $\begin{aligned} & \hline-0.0215^{* *} \\ & (0.0076) \end{aligned}$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2298 | 0.2510 | 0.2827 | 0.1547 | 0.1572 | 0.1611 |
| F | 129.386 | 123.872 | 129.875 | 1,269.505 | 849.454 | 765.046 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 18.16 | 20.18 | 21.05 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 35.13 | 36.38 | 37.99 |  |  |  |
| Sargan test chi2 | 1.89 | 1.78 | 1.67 |  |  |  |
| Total wealth amount (1,000PPP)(*) |  |  |  |  |  |  |
| Biased expectations indicator $>0$ | 0.2386*** | 0.2235** | 0.2145** | 0.0135** | 0.0128** | 0.0128* |
|  | (0.0938) | (0.0948) | (0.0942) | (0.0053) | (0.0052) | (0.0052) |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2813 | 0.3033 | 0.3057 | 0.1041 | 0.1086 | 0.1098 |
| F | 90.151 | 90.37 | 89.56 | 122.8 | 8.8 | 116.445 |



M1 includes optimistic indicator (after first-stage regression), gender, age, age squared, level of education (no education, primary education, lower secondary education, higher secondary education, post-secondary non-tertiary education and tertiary education (omitted)), marital status (married/cohabiting, separated/divorced, single and widow (omitted)), country fixed effects and wave fixed effects. M2 includes the same explanatory variables than M1, and additionally, number of days stayed at hospital during last year (surgery or medical tests), number of days stayed at hospital during last year (mental health), number of days stayed at other institutions during last year and number of visits to general practitioner during last year. M3 includes the same explanatory variables than M2, an additionally, relation with economic activity (employed, unemployed, retired, houseworking and other situations (omitted)), living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in 1,000 PPP). Bootstrap with 1,000 replications Robust standard errors between parenthesis. * statistically significant at $10 \%$; ** statistically significant at $5 \%$; *** statistically significant at $1 \%$. (*) Wealth is not included as explanatory variable in this regression.

Table B16. Effect of biased survival expectations (LO) and biased meteorological expectation (ME) over the amount of financial risk that the respondent is willing to take when making investments

|  | Using biased survival expectations <br> (IV- First-stage Table D.1) |  |  | Using meterological expectations (OLS) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | M1 | M2 | M3 | M1 | M2 | M3 |
| Take substantial risk expecting to earn substantial returns |  |  |  |  |  |  |
| Biased expectations indicator $>0$ | $\begin{aligned} & \text { 0.0277*** } \\ & (0.0035) \end{aligned}$ | $\begin{aligned} & 0.0251^{* * *} \\ & (0.0044) \end{aligned}$ | $\begin{aligned} & 0.0248^{* * *} \\ & (0.0022) \end{aligned}$ | $\begin{aligned} & 0.0060^{* * *} \\ & (0.0019) \end{aligned}$ | $\begin{aligned} & 0.0059 * * * \\ & (0.0019) \end{aligned}$ | $\begin{aligned} & 0.0059 * * * \\ & (0.0019) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2127 | 0.2153 | 0.2146 | 0.0782 | 0.0788 | 0.0791 |
| F | 40.102 | 40.371 | 38.379 | 191.685 | 134.615 | 121.816 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 30.75 | 32.44 | 33.70 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 24.14 | 25.70 | 24.18 |  |  |  |
| Sargan test chi2 | 1.16 | 1.09 | 0.91 |  |  |  |
| Take substantial risk expecting to earn substantial returns |  |  |  |  |  |  |
| Biased expectations indicator < 0 | $\begin{aligned} & \hline-0.0387 * * * \\ & (0.0021) \end{aligned}$ | $\begin{aligned} & \hline-0.0302 * * * \\ & (0.0020) \end{aligned}$ | $\begin{aligned} & -0.0282^{* * *} \\ & (0.0079) \end{aligned}$ | $\begin{aligned} & \hline-0.0161^{* * *} \\ & (0.0034) \end{aligned}$ | $\begin{aligned} & \hline-0.0155^{* * *} \\ & (0.0034) \end{aligned}$ | $\begin{aligned} & \hline-0.0153^{* * *} \\ & (0.0034) \end{aligned}$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2486 | 0.2477 | 0.2487 | 0.1410 | 0.1415 | 0.1424 |
| F | 126.595 | 122.422 | 116.730 | 991.257 | 659.660 | 590.479 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 28.76 | 29.98 | 30.16 |  |  |  |
| Durbin test chi2(1) /p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 33.78 | 34.89 | 34.09 |  |  |  |
| Sargan test chi2 | 1.45 | 1.51 | 1.34 |  |  |  |
| Take above average financial risks expecting to earn above average returns |  |  |  |  |  |  |
| Biased expectations indicator $>0$ | $\begin{aligned} & 1.1210^{* * *} \\ & (0.3544) \end{aligned}$ | $\begin{aligned} & 1.1031^{* * *} \\ & (0.4031) \end{aligned}$ | $\begin{aligned} & 1.0661^{* * *} \\ & (0.3880) \end{aligned}$ | $\begin{aligned} & 0.0319^{* * *} \\ & (0.0077) \end{aligned}$ | $\begin{aligned} & 0.0285^{* * *} \\ & (0.0077) \end{aligned}$ | $\begin{aligned} & 0.0277 * * * \\ & (0.0076) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.1619 | 0.1870 | 0.1844 | 0.1619 | 0.1650 | 0.1659 |
| F | 328.700 | 331.612 | 315.922 | 4,477.045 | 3,009.245 | 2,650.989 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 27.50 | 26.43 | 27.12 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 21.85 | 22.64 | 23.67 |  |  |  |
| Sargan test chi2 | 2.31 | 2.28 | 2.11 |  |  |  |
| Take above average financial risks expecting to earn above average returns |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | $-1.0475^{* * *}$ | $-0.9896 * *$ | $-0.9596 * *$ | $-0.0579 * * *$ | $-0.0540^{* * *}$ | $-0.0530^{* * *}$ |
|  | (0.3583) | (0.4037) | (0.3897) | (0.0080) | (0.0080) | (0.0080) |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2888 | 0.3176 | 0.3294 | 0.5174 | 0.5190 | 0.5198 |


| F | 1,680.102 | 1,722.838 | 1,606.077 | 24,852.998 | 16,429.099 | 14,421.279 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| Durbin test chi2(1)/p-value | $26.78$ | $28.14$ | $25.33$ |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 22.08 | 23.35 | 26.87 |  |  |  |
| Sargan test chi2 | 2.67 | 2.80 | 2.78 |  |  |  |
| Take average financial risks expecting to earn average returns |  |  |  |  |  |  |
| Biased expectations indicator $>0$ | $\begin{aligned} & -0.0771^{* * *} \\ & (0.0211) \end{aligned}$ | $\begin{aligned} & -0.0685^{* * *} \\ & (0.0208) \end{aligned}$ | $\begin{aligned} & -0.0673 * * * \\ & (0.0205) \end{aligned}$ | $\begin{aligned} & -0.0043^{* *} \\ & (0.0021) \end{aligned}$ | $\begin{aligned} & \hline-0.0042^{*} \\ & (0.0021) \end{aligned}$ | $\begin{aligned} & \hline-0.0040^{*} \\ & (0.0021) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.3052 | 0.3054 | 0.3066 | 0.02170 | 0.2182 | 0.2183 |
| F | 15.947 | 16.112 | 15.938 | 48.553 | 37.571 | 33.295 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 29.14 | 28.47 | 27.03 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 23.58 | 21.41 | 22.90 |  |  |  |
| Sargan test chi2 | 1.75 | 1.57 | 1.41 |  |  |  |
| Take average financial risks expecting to earn average returns |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | $\begin{aligned} & \hline 0.1080^{* * *} \\ & (0.0065) \end{aligned}$ | $\begin{aligned} & \hline 0.1060 * * * \\ & (0.0083) \end{aligned}$ | $\begin{aligned} & \begin{array}{l} 0.1015^{* * *} \\ (0.0017) \end{array} \end{aligned}$ | $\begin{aligned} & 0.0130^{* *} \\ & (0.0060) \end{aligned}$ | $\begin{aligned} & 0.0129 * * \\ & (0.0050) \end{aligned}$ | $\begin{aligned} & 0.0125^{* *} \\ & (0.0050) \end{aligned}$ |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| F | 0.3271 | 0.3274 | 0.3258 | 0.2305 | 0.2320 | 0.2327 |
| p | 36.641 | 35.324 | 34.572 | 217.999 | 150.787 | 134.715 |
| $\mathrm{R}^{2}$ | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 32.14 | 33.79 | 35.51 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 23.78 | 22.02 | 21.98 |  |  |  |
| Sargan test chi2 | 1.80 | 1.62 | 1.48 |  |  |  |
| Not willing to take any financial risks |  |  |  |  |  |  |
| Biased expectations indicator $>0$ | $\begin{aligned} & \hline-0.6822^{* * *} \\ & (0.1410) \end{aligned}$ | $\begin{aligned} & \hline-0.5752^{* * *} \\ & (0.1648) \end{aligned}$ | $\begin{aligned} & \hline-0.5598^{* * *} \\ & (0.1724) \end{aligned}$ | $\begin{aligned} & \hline-0.0116^{* *} \\ & (0.0048) \end{aligned}$ | $\begin{aligned} & \hline-0.0114^{* *} \\ & (0.0047) \end{aligned}$ | $\begin{aligned} & -0.0107^{* *} \\ & (0.0047) \end{aligned}$ |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| F | 0.1724 | 0.2097 | 0.2147 | 0.1608 | 0.1654 | 0.1676 |
| P | 158.385 | 155.737 | 152.498 | 1.329.275 | 902.858 | 802.377 |
| $\mathrm{R}^{2}$ | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 43.78 | 42.53 | 45.87 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 28.02 | 26.43 | 25.51 |  |  |  |
| Sargan test chi2 | 0.64 | 0.58 | 0.53 |  |  |  |
| Not willing to take any financial risks |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | $0.4028^{* * *}$ | 0.3802*** | 0.3507*** | $0.0176^{* *}$ | $0.0172^{* *}$ | 0.0166** |
|  | (0.0492) | (0.0783) | (0.0861) | (0.0079) | (0.0078) | (0.0078) |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| F | 0.6005 | 0.6085 | 0.6109 | 0.5709 | 0.5735 | 0.5747 |
| P | 1,255.924 | 1,184.734 | 1,155.490 | 9,229.947 | 6,126.153 | 5,384.768 |
| $\mathrm{R}^{2}$ | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | 41.36 | 41.79 | 43.80 |  |  |  |
| Durbin test chi2(1)/p-value | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 24.75 | 25.78 | 24.89 |  |  |  |
| Sargan test chi2 | 0.60 | 0.55 | 0.31 |  |  |  |
| Education, marital status <br> Days at hospital for surgery/tests, due to mental problems, days at other institutions, visits to GP Relation with economic activity, lives alone, lives in rural area, income and wealth <br> Country fixed effects <br> Time fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |
|  | No | Yes | Yes | No | Yes | Yes |
|  | No | No | Yes | No | No | Yes |
|  | Yes | Yes | Yes | Yes | Yes | Yes |
|  | Yes | Yes | Yes | Yes | Yes | Yes |

M1 includes optimistic indicator (after first-stage regression), gender, age, age squared, level of education (no education, primary education, lower secondary education, higher secondary education, post-secondary non-tertiary education and tertiary education (omitted)), marital status (married/cohabiting, separated/divorced, single and widow (omitted)), country fixed effects and wave fixed effects. M2 includes the same explanatory variables than M1, and additionally, number of days stayed at hospital during last year (surgery or medical tests), number of days stayed at hospital during last year (mental health), number of days stayed at other institutions during last year and number of visits to general practitioner during last year. M3 includes the same explanatory variables than M2, an additionally, relation with economic activity (employed, unemployed, retired, houseworking and other situations (omitted)), living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in 1,000 PPP). Bootstrap with 1,000 replications Robust standard errors between parenthesis. * statistically significant at $10 \%$; ** statistically significant at $5 \%$; *** statistically significant at $1 \%$.

Table B17. Effect of biased survival expectations (LO) and biased meteorological expectation (ME) over social contacts

|  | Using biased survival expectations <br> (IV- First-stage Table D.1) |  |  | Using meterological expectations (OLS) |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | M1 | M2 | M3 | M1 | M2 | M3 |
| Social contacts |  |  |  |  |  |  |
| Biased expectations indicator $>0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $0.2814^{* * *}$ $(0.0788)$ 50,058 0.6525 $1,457.312$ 0.0000 27.05 $(0.0000)$ 23.56 1.17 | $0.2790^{* * *}$ $(0.0699)$ 50,058 0.6681 $1,436.108$ 0.0000 28.67 $(0.0000)$ 24.25 1.10 | $0.2509^{* * *}$ $(0.0613)$ 50,058 0.6662 $1,387.521$ 0.0000 25.13 $(0.0000)$ 24.60 0.98 | $\begin{aligned} & \hline 0.4052^{* *} \\ & (0.1506) \\ & 53,336 \\ & 0.6491 \\ & 1,831.377 \\ & 0.0000 \end{aligned}$ | $0.3136^{* *}$ $(0.1395)$ 53,336 0.6669 $1,828.466$ 0.0000 | $0.3085^{* *}$ $(0.1388)$ 53,336 0.6681 $1,780.234$ 0.0000 |
| Social contacts |  |  |  |  |  |  |
| Biased expectations indicator $<0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo’s test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $-0.0579^{* * *}$ $(0.0166)$ 19,842 0.4986 348.214 0.0000 44.24 $(0.0000)$ 29.80 1.67 | $-0.0565^{* * *}$ $(0.0121)$ 19,842 0.4043 328.724 0.0000 43.25 $(0.0000)$ 29.17 1.54 | $-0.0562^{* * *}$ $(0.0125)$ 19,842 0.4039 318.509 0.0000 43.19 $(0.0000)$ 29.04 1.49 | $-0.0136^{* *}$ $(0.0042)$ 15,979 0.3178 492.488 0.0000 | $-0.0122^{* * *}$ $(0.0042)$ 15,979 0.3214 461.813 0.0000 | $-0.0117^{* * *}$ $(0.0041)$ 15,979 0.3224 449.429 0.0000 |
| Feels things are out of control: often |  |  |  |  |  |  |
| Biased expectations indicator $>0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> p <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $-0.4981^{* *}$ $(0.2086)$ 50,058 0.1012 125.170 0.0000 22.15 $(0.0000)$ 25.50 1.48 | $-0.4892^{* *}$ $(0.1933)$ 50,058 0.1078 131.952 0.0000 21.41 $(0.0000)$ 25.01 1.37 | $-0.4635^{* *}$ $(0.2090)$ 50,058 0.1122 265.282 0.0000 24.01 $(0.0000)$ 26.48 1.22 | $-0.0195^{* * *}$ $(0.0055)$ 53,336 0.0361 $1,329.886$ 0.0000 | $-0.0187^{* * *}$ $(0.0055)$ 53,336 0.0381 918.184 0.0000 | $\begin{aligned} & \hline-0.0101^{*} \\ & (0.0053) \\ & 53,336 \\ & 0.0897 \\ & 1,500.742 \\ & 0.0000 \end{aligned}$ |
| Feels things are out of control: often |  |  |  |  |  |  |
| Biased expectations indicator $<0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> P <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $0.9489^{* *}$ $(0.3339)$ 19,842 0.1217 55.947 0.0000 20.18 $(0.0000)$ 23.44 1.59 | $-0.8213^{* *}$ <br> $(0.3195)$ <br> 19,842 <br> 0.1228 <br> 59.448 <br> 0.0000 <br> 21.05 <br> $(0.0000)$ <br> 23.78 <br> 1.45 | $0.7994^{* *}$ $(0.2984)$ 19,842 0.1362 128.050 0.0000 23.15 $(0.0000)$ 24.10 1.36 | $\begin{aligned} & \hline 0.0278 * * * \\ & (0.0051) \\ & 15,979 \\ & 0.0461 \\ & 513.889 \\ & 0.0000 \end{aligned}$ | $0.0275 * * *$ $(0.0051)$ 15,979 0.0504 368.423 0.0000 | $0.0245^{* * *}$ $(0.0045)$ 15,979 0.1237 642.820 0.0000 |
| Feels things are out of control: sometimes |  |  |  |  |  |  |
| Biased expectations indicator $>0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> P <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo’s test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $-0.5256^{* * *}$ $(0.1129)$ 50,058 0.1544 266.697 0.0000 23.58 $(0.0000)$ 22.71 1.59 | $-0.2848^{* *}$ $(0.1074)$ 50,058 0.1578 253.100 0.0000 23.47 $(0.0000)$ 22.18 1.44 | $-0.2255^{* *}$ $(0.1075)$ 50,058 0.1610 312.104 0.0000 23.60 $(0.0000)$ 23.05 1.29 | $-0.0777^{* * *}$ $(0.0089)$ 53,336 0.0767 $2,951.259$ 0.0000 | $-0.0763^{* * *}$ $(0.0089)$ 53,336 0.0778 $1,954.878$ 0.0000 | $-0.0633^{* * *}$ $(0.0088)$ 53,336 0.0972 $1,640.135$ 0.0000 |
| Feels things are out of control: sometimes |  |  |  |  |  |  |
| Biased expectations indicator $<0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> P <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $0.5291^{* * *}$ $(0.1695)$ 19,842 0.1605 120.351 0.0000 19.73 $(0.0000)$ 25.30 1.45 | $0.4425^{* *}$ $(0.1653)$ 19,842 0.1641 116.192 0.0000 19.81 $(0.0000)$ 25.65 1.23 | $0.4328^{* *}$ $(0.1522)$ 19,842 0.1718 117.113 0.0000 20.03 $(0.0000)$ 27.02 1.17 | $0.0503^{* *}$ $(0.0217)$ 15,979 0.0784 905.714 0.0000 | $0.0498^{* *}$ $(0.0217)$ 15,979 0.0791 595.760 0.0000 | 0.0481** $(0.0215)$ 15,979 0.0943 474.465 0.0000 |
| Feels things are out of control: rarely |  |  |  |  |  |  |
| Biased expectations indicator $>0$ <br> N <br> $\mathrm{R}^{2}$ <br> F <br> P <br> Durbin test chi2(1)/p-value <br> Stock \& Yogo's test (critical value at 5\%: 19.93) <br> Sargan test chi2 | $-0.2902^{* * *}$ $(0.0442)$ 50,058 0.1923 388.476 0.0000 19.12 $(0.0000)$ 35.11 1.31 | $-0.2788^{* * *}$ $(0.0349)$ 50,058 0.1975 369.710 0.0000 19.65 $(0.0000)$ 35.17 1.23 | $-0.1894^{* * *}$ $(0.0831)$ 50,058 0.1984 372.668 0.0000 19.71 $(0.0000)$ 36.23 1.11 | $\begin{aligned} & 0.0012 \\ & (0.0097) \\ & 53,336 \\ & 0.1086 \\ & 4,329.949 \\ & 0.0000 \end{aligned}$ | -0.0013 $(0.0097)$ 53,336 0.1102 $2,870.911$ 0.0000 | -0.0038 $(0.0097)$ 53,336 0.1150 $1,978.499$ 0.0000 |
| Feels things are out of control: rarely |  |  |  |  |  |  |
| Biased expectations indicator $<0$ N | $\begin{aligned} & \hline 1.1085^{* * *} \\ & (0.3191) \\ & 19,842 \end{aligned}$ | $\begin{aligned} & \hline 0.9103^{* *} \\ & (0.3018) \\ & 19,842 \end{aligned}$ | $\begin{aligned} & \hline 0.6532^{* *} \\ & (0.3274) \\ & 19,842 \end{aligned}$ | $\begin{aligned} & \hline 0.0367 \\ & (0.0225) \\ & 15,979 \end{aligned}$ | $\begin{aligned} & \hline 0.0352 \\ & (0.0225) \\ & 15,979 \end{aligned}$ | $\begin{aligned} & \hline 0.0317 \\ & (0.0224) \\ & 15,979 \end{aligned}$ |


| $\mathrm{R}^{2}$ | 0.2016 | 0.2057 | 0.2091 | 0.0950 | 0.0966 | 0.1064 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| F | 128.663 | 130.370 | 135.482 | 1,117.617 | 741.814 | 542.637 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| Durbin test chi2(1)/p-value | $\begin{aligned} & 23.41 \\ & (0.0000) \end{aligned}$ | $\begin{aligned} & 23.50 \\ & (0.0000) \end{aligned}$ | $\begin{aligned} & 24.65 \\ & (0.0000) \end{aligned}$ |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 34.34 | 34.19 | 34.76 |  |  |  |
| Sargan test chi2 | 1.17 | 1.24 | 1.21 |  |  |  |
| Feels things are out of control: never |  |  |  |  |  |  |
| Biased expectations indicator $>0$ | 0.6580*** | 0.4500** | 0.4430** | 0.0946*** | 0.0948*** | 0.0755*** |
|  | (0.1656) | (0.1487) | (0.1405) | (0.0098) | (0.0098) | (0.0096) |
| N | 50,058 | 50,058 | 50,058 | 53,336 | 53,336 | 53,336 |
| $\mathrm{R}^{2}$ | 0.2105 | 0.2133 | 0.2178 | 0.1322 | 0.1335 | 0.1650 |
| F | 565.285 | 546.860 | 629.164 | 5,414.361 | 3,572.653 | 3,009.816 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| Durbin test chi2(1)/p-value | 33.57 | 33.80 | 34.81 |  |  |  |
|  | (0.0000) | (0.0000) | (0.0000) |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | 23.61 | 23.90 | 23.98 |  |  |  |
| Sargan test chi2 | 1.45 | 1.67 | 1.89 |  |  |  |
| Feels things are out of control: never |  |  |  |  |  |  |
| Biased expectations indicator $<0$ | -0.3644*** | -0.3335*** | ${ }^{-0.2608 * * *}$ | -0.0586** | -0.0570** | -0.0551** |
|  | (0.0510) | (0.0493) | (0.0480) | (0.0230) | (0.0229) | (0.0225) |
| N | 19,842 | 19,842 | 19,842 | 15,979 | 15,979 | 15,979 |
| $\mathrm{R}^{2}$ | 0.2256 | 0.2279 | 0.2380 | 0.1073 | 0.1095 | 0.1457 |
| F | 119.510 | 114.634 | 131.859 | 1,279.643 | 852.819 | 777.083 |
| P | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
| Durbin test chi2(1)/p-value | $31.11$ | $31.15$ | $31.43$ |  |  |  |
| Stock \& Yogo's test (critical value at 5\%: 19.93) | (0.0000) 24.78 | $24.87$ | $24.50$ |  |  |  |
| Sargan test chi2 | 1.89 | 1.69 | 1.82 |  |  |  |
| Education, marital status | Yes | Yes | Yes | Yes | Yes | Yes |
| Education, marital status | Yes | Yes | Yes | Yes | Yes | Yes |
| Days at hospital for surgery/tests, due to mental problems, days at other institutions, visits to GP | No | Yes | Yes | No | Yes | Yes |
| Relation with economic activity, lives alone, lives in rural area, income and wealth | No | No | Yes | No | No | Yes |
| Country fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |
| Time fixed effects | Yes | Yes | Yes | Yes | Yes | Yes |

M1 includes optimistic indicator (after first-stage regression), gender, age, age squared, level of education (no education, primary education, lower secondary education, higher secondary education, post-secondary non-tertiary education and tertiary education (omitted), marital status (married/cohabiting, separated/divorced, single and widow (omitted)), country fixed effects and wave fixed effects. M2 includes the same explanatory variables than M1, and additionally, number of days stayed at hospital during last year (surgery or medical tests), number of days stayed at hospital during last year (mental health), number of days stayed at other institutions during last year and number of visits to general practitioner during last year. M3 includes the same explanatory variables than M2, an additionally, relation with economic activity (employed, unemployed, retired, houseworking and other situations (omitted)), living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in 1,000 PPP).

Table B18. Effect of one standard deviation increase (or decrease) of biased survival expectations (LO) and biased meteorological expectation (ME) over social contacts

|  | $\begin{gathered} \text { One standard } \\ \text { deviation } \\ \text { increase of LO>0 } \\ \hline \end{gathered}$ | $\begin{gathered} \text { One standard } \\ \text { deviation } \\ \text { increase of } \mathrm{LO}<0 \end{gathered}$ | One standard deviation increase of ME $>0$ | One standard deviation increase of ME $<0$ |
| :---: | :---: | :---: | :---: | :---: |
| Social contacts Effect in percentual points With respect to the mean probability | $\begin{array}{r} 6.0873 \mathrm{pp} \\ 38.7234 \% \\ \hline \end{array}$ | $\begin{aligned} & -1.0060 \mathrm{pp} \\ & -6.3994 \% \\ & \hline \end{aligned}$ | $\begin{array}{r} 5.8276 \mathrm{pp} \\ 37.0710 \% \\ \hline \end{array}$ | $\begin{aligned} & -0.1719 \mathrm{pp} \\ & -1.0933 \% \\ & \hline \end{aligned}$ |
| Feels things are out of control: often Effect in percentual points With respect to the mean probability | $\begin{array}{r} -11.2445 \mathrm{pp} \\ -152.5714 \% \\ \hline \end{array}$ | $\begin{gathered} 14.3093 \mathrm{pp} \\ 194.1555 \% \end{gathered}$ | $\begin{array}{r} -0.1908 \mathrm{pp} \\ -2.5887 \% \\ \hline \end{array}$ | $\begin{gathered} 0.3599 \text { pp } \\ 4.8834 \% \end{gathered}$ |
| Feels things are out of control: sometimes <br> Effect in percentual points <br> With respect to the mean probability | $\begin{array}{r} -5.4706 \mathrm{pp} \\ -28.2866 \% \\ \hline \end{array}$ | $\begin{array}{r} 7.7471 \mathrm{pp} \\ 40.0575 \% \\ \hline \end{array}$ | $\begin{aligned} & -1.1957 \mathrm{pp} \\ & -6.1827 \% \\ & \hline \end{aligned}$ | $\begin{aligned} & 0.7066 \mathrm{pp} \\ & 3.6535 \% \\ & \hline \end{aligned}$ |
| Feels things are out of control: rarely <br> Effect in percentual points <br> With respect to the mean probability | $\begin{array}{r} -4.5948 \mathrm{pp} \\ -21.4312 \% \\ \hline \end{array}$ | $\begin{aligned} & 11.6923 \mathrm{pp} \\ & 54.5349 \% \\ & \hline \end{aligned}$ | No signif | No signif |
| Feels things are out of control: never <br> Effect in percentual points <br> With respect to the mean probability | $\begin{aligned} & 10.7472 \mathrm{pp} \\ & 42.3618 \% \end{aligned}$ | $\begin{gathered} -4.6683 \mathrm{pp} \\ -18.4009 \% \end{gathered}$ | $\begin{gathered} 1.4262 \mathrm{pp} \\ 5.6216 \text { \% } \end{gathered}$ | $\begin{aligned} & -0.8094 \mathrm{pp} \\ & -3.1905 \% \end{aligned}$ |

The effects for LO are obtained after estimating equation (4) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B17 (and on Table D. 1 for first-stage regression). The effects for ME are obtained after estimating equation (8) with the explanatory variables included in specification M3. Estimated coefficients are reported on Table B17. Mean values for the dependent variables are reported on Table B6.

## Appendix C. Discrete time-hazard model

Estimating biased survival expectations. The departure point of our analysis is the duration model of Jenkins (1995). He estimated a discrete-time hazard model to explain the effect of different covariates over the transition to retirement. The key point of his model is to use as unit of analysis the time of risk of an event, so that time periods prior to selection into the sample can be ignored. In this way, the estimation of the discrete-time hazard model is simplified to the estimation of a binary outcome (retired vs. not retired in the case of Jenkins (1995), or live vs. died, in our case). We create our sample of interest selecting alive individuals at each wave of SHARE following them through the next waves, until they are observed to decease or they are censored. Transition to death is estimated using a discrete-time hazard model which allows us to analyze the impact of socioeconomic characteristic and health indicators over the probability of not surviving.

To understand how the stock-sampling works, consider first individuals who are alive at wave 1 . In the following waves, they may be still alive or have died. Using Jenkins' (1995) notation $t=\tau$ is the first observation of the stock sample, that is, the first period at which the individual is at risk of decease. By the time of the second interview (for the subsample of follow-up respondents), each individual may be still living (censored duration data $\delta_{i}=0$ ) or have died (complete duration data $\delta_{i}=1$ ). $t=\tau+s_{i}$ is the year of the decease if $\delta_{i}=0$, and is the final year of our observation period if $\delta_{i}=0$. Therefore, each respondent $i$ contributes to $s_{i}$ of life spell data. Given that the probability of dying at each time period $t$ provides information on the duration distribution, the discrete-time hazard $\left(h_{i t}\right)$ is defined as:

$$
\begin{equation*}
h_{i t}=P\left[T_{i}=t \mid T_{i} \geq t, \Omega_{i t}\right] \tag{C.1}
\end{equation*}
$$

where $\Omega_{i t}$ is a vector of explanatory variables ${ }^{52}$ that varies with time and $T_{i}$ is a random variable representing the time at which the decease is observed. Then, upon the restriction of being alive at the beginning of the spell, the conditional probability of observing an uncompleted spell (i.e., being alive at the next interview) is the following:

$$
\begin{equation*}
\operatorname{Pr}\left[T_{i}>t \mid T_{i}>\tau-1\right]=\prod_{t=\tau}^{\tau+s_{i}}\left(1-h_{i t}\right) \tag{C.2}
\end{equation*}
$$

The probability of dying between the initial observation ( $\tau$ ) and next interview is:

[^29]\[

$$
\begin{equation*}
\operatorname{Pr}\left[T_{i}>t \mid T_{i}=\tau-1\right]=h_{i \tau+s_{i}} \prod_{t=\tau}^{\tau+s_{i}-1}\left(1-h_{i t}\right)=\frac{h_{i \tau+s_{i}}}{1-h_{i \tau+s_{i}}} \prod_{t=\tau}^{\tau+s_{i}}\left(1-h_{i t}\right) \tag{C.3}
\end{equation*}
$$

\]

And the log-likelihood function for the complete sample is:

$$
\begin{equation*}
\log L=\sum_{i=1}^{N} \delta_{i} \log \left(\frac{h_{i \tau+s_{i}}}{1-h_{i \tau+s_{i}}}\right)+\sum_{i=1}^{N} \sum_{t=\tau}^{\tau+s_{i}} \log \left(1-h_{i t}\right) \tag{C.4}
\end{equation*}
$$

As we only observe individuals who are alive by the time of the first interview, we can simplify the loglikelihood by defining a binary indicator. For those who are alive at the end of the observation period: $y_{i t}=0$, for all periods. For those who die: $y_{i t}=0$ for all periods, except the last one in which case $y_{i t}=1$. It can also be expressed as: $y_{i t}=1$ if $t=\tau+s_{i}$ and $\delta_{i}=1$ (for decease-observation) and $y_{i t}=0$ otherwise (for aliveobservation). After introducing this binary indicator in the log-likelihood function we get:

$$
\begin{equation*}
\log L=\sum_{i=1}^{N} \sum_{t=\tau}^{\tau+s_{i}} y_{i} \log \left(\frac{h_{i \tau+s_{i}}}{1-h_{i \tau+s_{i}}}\right)+\sum_{i=1}^{N} \sum_{t=\tau}^{\tau+s_{i}} \log \left(1-h_{i t}\right) \tag{C.5}
\end{equation*}
$$

Consequently, (C.5) seems like the log-likelihood function of a binary variable, where the unit of analysis is the spell period.

In the estimation of the discrete time hazard model, we have adjusted five different functional forms: exponential, Gompertz, Weibull, log-normal and log-logistic. To determine which one fits better to the data we have obtained the cox-snell residuals and two information criteria (AIC y BIC). Figure C1 shows the cox-snell residuals for all the functional forms. As the model fits better to the data, the cox-snell residuals seem like an exponential distribution with scale parameter equal to 1 . For this purpose, we have estimated the Kaplan-Meier cumulative hazard function using the cox-snell residuals as the time variable and plotted them against the coxsnell residuals. We observe that the cox-snell residuals are nearer to the bisector of the first quadrant for the Weibull function. Additionally, the AIC and BIC attain the lowest values for the Weibull model (Table C1).

Figure C1. Cox-snell residuals of the survival models


All models include the same explanatory variables: age age-squared, gender, level of education, obese, overweight, chronic illness (high blood pressure, cancer, stroke, diabetes, heart attack, lung disease, Alzheimer), number of days stayed at hospital during last year (not due to mental health problems), number of days stayed in hospital due to mental health problems, number of days stayed in other institutions during last year, number of visits to general practitioner during last year, feeling depressed, drinking alcohol every day or 5-6 days per week, sedentary lifestyle, smokes now, has smoked before (not now), being employed, living alone, living in rural area/village/small town, adjusted income, adjusted wealth, country fixed effects, time fixed effects..

Table C1. Comparison of AIC and BIC statistics

|  | Gompertz | Weibull | Log-normal | Log-logistic | Exponential |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Baseline model (N=69,900) |  |  |  |  |  |
| AIC | $64,248.669$ | $61,098.945$ | $65,622.216$ | $65,032.539$ | $61,974.631$ |
| BIC | $64,249.044$ | $61,099.320$ | $65,622.591$ | $65,032.915$ |  |
| Excluding focal responses |  |  |  |  |  |
| AIC | $67,462.102$ | $64,154.892$ | $68,904.327$ | $68,285.166$ | $65,074.363$ |
| BIC | $67,462.496$ | $64,155.286$ | $68,904.721$ | $68,285.561$ | $65,074.747$ |
| Excluding respondents with erroneous responses to probabilistic question |  |  |  |  |  |
| AIC | $72,185.450$ | $68,646.735$ | $73,728.630$ | $73,066.128$ | $69,630.568$ |
| BIC | $72,185.871$ | $68,647.156$ | $73,729.051$ | $73,066.550$ | $69,630.979$ |
| Including parental characteristics as explanatory variables in the baseline model |  |  |  |  |  |
| AIC | $60,120.778$ | $57,365.864$ | $61,315.941$ | $60,803.308$ | $58,133.776$ |
| BIC | $60,121.104$ | $57,366.193$ | $61,316.267$ | $60,803.635$ | $58,134.097$ |

All models include the same explanatory variables: age age-squared, gender, level of education, obese, overweight, chronic illness (high blood pressure, cancer, stroke, diabetes, heart attack, lung disease, Alzheimer), number of days stayed at hospital during last year (not due to mental health problems), number of days stayed in hospital due to mental health problems, number of days stayed in other institutions during last year, number of visits to general practitioner during last year, feeling depressed, drinking alcohol every day or 5-6 days per week, sedentary lifestyle, smokes now, has smoked before (not now), being employed, living alone, living in rural area/village/small town, adjusted income, adjusted wealth, country fixed effects, time fixed effects..

Table C2. Comparison of subjective survival probability and objective survival probability gender

|  | N | $\begin{aligned} & \text { SSP } \\ & \text { (\%) } \end{aligned}$ | $\begin{aligned} & \hline \text { OSP } \\ & \text { (\%) } \end{aligned}$ | LO Difference SSP-OSP | Optimistic SSP>OSP (\%) | Pesimistic SSP<OSP <br> (\%) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Deceased |  |  |  |  |  |  |
| Men | 1,142 | $\begin{gathered} 43.01 \\ (33.75) \end{gathered}$ | $\begin{gathered} 16.17 \\ (23.10) \end{gathered}$ | $\begin{gathered} 26.84 \\ (34.42) \end{gathered}$ | 74.33 | 25.67 |
| Women | 898 | $\begin{gathered} 37.96 \\ (31.65) \end{gathered}$ | $\begin{aligned} & 18.33 \\ & 26.03) \end{aligned}$ | $\begin{gathered} 19.63 \\ (34.89) \end{gathered}$ | 65.24 | 34.76 |
| All | 2,040 | $\begin{gathered} 40.63 \\ (32.87) \end{gathered}$ | $\begin{gathered} 17.18 \\ (24.54) \end{gathered}$ | $\begin{gathered} 23.44 \\ (34.82) \\ \hline \end{gathered}$ | 70.04 | 29.96 |
| Survivor |  |  |  |  |  |  |
| Men | 30,829 | $\begin{gathered} 64.51 \\ (28.97) \end{gathered}$ | $\begin{gathered} \hline 42.51 \\ (28.15) \end{gathered}$ | $\begin{gathered} 22.00 \\ (35.03) \end{gathered}$ | 72.89 | 27.11 |
| Women | 37,031 | $\begin{gathered} 62.10 \\ (29.87) \end{gathered}$ | $\begin{gathered} 52.67 \\ (29.55) \end{gathered}$ | $\begin{gathered} 9.43 \\ (33.85) \end{gathered}$ | 60.90 | 39.10 |
| All | 67,860 | $\begin{gathered} 63.21 \\ (29.48) \\ \hline \end{gathered}$ | $\begin{gathered} 47.99 \\ (29.35) \\ \hline \end{gathered}$ | $\begin{gathered} 15.21 \\ (34.96) \end{gathered}$ | 63.41 | 36.59 |

SSP: subjective survival probability. OSP: objective survival probability
Standard errors between parenthesis. Using calibrated sampling weights

Table C3. Comparison of subjective survival probability and objective survival probability by country


SSP: subjective survival probability. OSP: objective survival probability
Standard errors between parenthesis. Using calibrated sampling weights
Group I: Italy and Spain. Group II: Belgium and France. Group III: Austria and Denmark. Group IV: Germany, Sweden and Switzerland.
Survivors subsample (by country)

|  | N | $\begin{aligned} & \hline \text { SSP } \\ & \text { (\%) } \end{aligned}$ | $\begin{gathered} \hline \text { OSP } \\ \text { (\%) } \end{gathered}$ | LO Difference SSP-OSP | $\begin{gathered} \hline \text { Optimistic } \\ \text { SSP>OSP } \\ (\%) \\ \hline \end{gathered}$ | $\begin{gathered} \hline \text { Pesimistic } \\ \text { SSP }<\text { OSP } \\ (\%) \\ \hline \end{gathered}$ |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Austria | 7,680 | $\begin{gathered} \hline 64.48 \\ (29.36) \end{gathered}$ | $\begin{gathered} \hline 45.24 \\ (28.32) \end{gathered}$ | $\begin{gathered} 19.24 \\ (33.13) \end{gathered}$ | 71.23 | 28.77 |
| Belgium | 10,531 | $\begin{gathered} 59.76 \\ (27.42) \end{gathered}$ | $\begin{gathered} 50.81 \\ (29.27) \end{gathered}$ | $\begin{gathered} 8.96 \\ (32.67) \end{gathered}$ | 59.18 | 40.82 |
| Denmark | 6,227 | $\begin{gathered} 71.26 \\ (27.89) \end{gathered}$ | $\begin{gathered} 46.98 \\ (28.45) \end{gathered}$ | $\begin{gathered} 24.28 \\ (32.19) \end{gathered}$ | 77.92 | 22.08 |
| France | 8,231 | $\begin{gathered} 61.48 \\ (28.99) \end{gathered}$ | $\begin{gathered} 51.94 \\ (29.28) \end{gathered}$ | $\begin{gathered} 9.53 \\ (33.64) \end{gathered}$ | 60.38 | 39.62 |
| Germany | 6,997 | $\begin{gathered} 60.91 \\ (29.80) \end{gathered}$ | $\begin{gathered} 51.71 \\ (28.94) \end{gathered}$ | $\begin{gathered} 9.20 \\ (34.04) \end{gathered}$ | 60.85 | 39.15 |
| Italy | 7,516 | $\begin{gathered} 65.56 \\ (30.13) \end{gathered}$ | $\begin{gathered} 46.29 \\ (29.38) \end{gathered}$ | $\begin{gathered} 19.27 \\ (36.40) \end{gathered}$ | 70.29 | 29.71 |
| Spain | 7,751 | $\begin{gathered} 64.57 \\ (29.19) \end{gathered}$ | $\begin{gathered} 38.39 \\ (28.26) \end{gathered}$ | $\begin{gathered} 26.19 \\ (34.42) \end{gathered}$ | 76.82 | 23.18 |
| Sweden | 6,921 | $\begin{gathered} 63.62 \\ (30.48) \end{gathered}$ | $\begin{gathered} 43.92 \\ (29.13) \end{gathered}$ | $\begin{gathered} 19.71 \\ (33.26) \end{gathered}$ | 71.25 | 28.74 |
| Switzerland | 6,006 | $\begin{gathered} 68.05 \\ (26.66) \\ \hline \end{gathered}$ | $\begin{gathered} 51.68 \\ (27.89) \\ \hline \end{gathered}$ | $\begin{gathered} 16.37 \\ (32.65) \\ \hline \end{gathered}$ | 68.84 | 31.16 |
| Group I | 15,267 | $\begin{gathered} 65.17 \\ (29.76) \end{gathered}$ | $\begin{gathered} 43.15 \\ (29.20) \end{gathered}$ | $\begin{gathered} 22.02 \\ (35.79) \end{gathered}$ | 72.88 | 27.12 |
| Group II | 18,762 | $\begin{gathered} 61.19 \\ (28.74) \end{gathered}$ | $\begin{gathered} 51.75 \\ (29.28) \end{gathered}$ | $\begin{gathered} 9.44 \\ (33.48) \end{gathered}$ | 60.18 | 39.82 |
| Group III | 13,907 | $\begin{gathered} 67.00 \\ (29.01) \end{gathered}$ | $\begin{gathered} 45.88 \\ (28.38) \end{gathered}$ | $\begin{gathered} 21.11 \\ (32.87) \end{gathered}$ | 73.71 | 26.29 |
| Group IV | 19,924 | $\begin{gathered} 61.93 \\ (29.64) \\ \hline \end{gathered}$ | $\begin{gathered} 50.96 \\ (28.94) \\ \hline \end{gathered}$ | $\begin{gathered} 10.97 \\ (34.01) \\ \hline \end{gathered}$ | 62.70 | 37.30 |
| All countries | 67,860 | $\begin{gathered} 63.21 \\ (29.48) \\ \hline \end{gathered}$ | $\begin{gathered} 47.99 \\ (29.35) \\ \hline \end{gathered}$ | $\begin{gathered} 15.21 \\ (34.96) \\ \hline \end{gathered}$ | 63.41 | 36.59 |

SSP: subjective survival probability. OSP: objective survival probability
Standard errors between parenthesis. Using calibrated sampling weights
Group I: Italy and Spain. Group II: Belgium and France. Group III: Austria and Denmark. Group IV: Germany, Sweden and Switzerland.

Appendix D. First-stage regressions and robustness checks
Figure D1. Distribution of subjective survival probabilities


Note: 2,040 observations for the deceased subsample; 67,860 observations for the survivors subsample.

Figure D2. Kernel density function for the "biased survival indicator" (difference between subjective survival expectations and objective survival expectations): total sample vs. sample excluding focal responses


Including focal responses: $\mathrm{N}=69,900$. Excluding focal responses ( $0,50,100$ ): $\mathrm{N}=40,893$.

Figure D3. Kernel density function for the "biased survival indicator" (difference between subjective survival expectations and objective survival expectations): total sample vs. sample excluding erroneous responses to probabilistic question.


Probabilistic question: "If the chance of getting a disease is 10 per cent, how many people out of 1000 (one thousand) would be expected to get the disease?" Total sample: N=69,900. Excluding erroneous responses to probabilistic question: N=30,183.

Figure D4. Comparison between observed bias, predicted bias using SHARELIFE respondents and all instruments and predicted bias using SHARELIFE respondents and only 2 instruments (blood test and blood pressure).


Note: Black straight line corresponds to observed LO (difference between objective and subjective survival expectations represented on Figure 1). Black line with squares corresponds to predicted LO using only SHARELIFE respondents and all instruments (parent's age of deceased, blood test and blood pressure) ( $4^{\text {th }}$ column of Table D1). Parent`s age of decease has been imputed for those fathers/mothers alive at the time of the survey. Grey straight line corresponds to predicted LO using only SHARELIFE respondents, but using only two instruments (blood test and blood pressure) ( $5^{\text {th }}$ column of Table D1).

Figure D5. Comparison between observed bias, predicted bias and predicted bias taking the average for each respondent.


Note: Black straight line corresponds to observed LO (difference between objective and subjective survival expectations represented on Figure 1). Black line with squares corresponds to predicted LO using the whole sample of respondents (Table D1). Grey straight line corresponds to predicted LO using the whole sample of respondents (Table D1) and taking the mean for each individual.

Table D1. First stage regressions for the "biased survival" indicator (difference between SSP and OSP)

|  | Using the whole sample of respondents |  |  | Sample of respondents who have answered SHARELIFE |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | M1 | M2 | M3 | M4 | M5 |
| Mother's age at decease ${ }^{\text {a }}$ | $\begin{aligned} & \hline 0.0011^{* * *} \\ & (0.0001) \end{aligned}$ | $\begin{aligned} & 0.0012 * * * \\ & (0.0001) \end{aligned}$ | $\begin{aligned} & \text { 0.0012*** } \\ & (0.0001) \end{aligned}$ | $\begin{aligned} & \hline 0.0010^{* * *} \\ & (0.0001) \end{aligned}$ |  |
| Father's age at decease ${ }^{\text {b }}$ | $\begin{aligned} & 0.0015^{* * *} \\ & (0.0001) \end{aligned}$ | $\begin{aligned} & 0.0017^{* * *} \\ & (0.0001) \end{aligned}$ | $\begin{aligned} & 0.0017^{* * *} \\ & (0.0001) \end{aligned}$ | $\begin{aligned} & 0.0014^{* * *} \\ & (0.0001) \end{aligned}$ |  |
| Mother alive | $\begin{aligned} & 0.0111^{* *} \\ & (0.0051) \end{aligned}$ | $\begin{aligned} & 0.0114^{* *} \\ & (0.0050) \end{aligned}$ | $\begin{aligned} & 0.0114 * * \\ & (0.0050) \end{aligned}$ | $\begin{aligned} & 0.0107 * * \\ & (0.0048) \end{aligned}$ |  |
| Father alive | $\begin{aligned} & 0.0242^{* * *} \\ & (0.0047) \end{aligned}$ | $\begin{aligned} & 0.0248^{* * *} \\ & (0.0046) \end{aligned}$ | $\begin{aligned} & 0.0248^{* * *} \\ & (0.0046) \end{aligned}$ | $\begin{aligned} & 0.0235^{* * *} \\ & (0.0043) \end{aligned}$ | - |
| Blood pressure checked last year | - | - | - | $\begin{aligned} & -11.6010^{* * *} \\ & (0.0034) \end{aligned}$ | $\begin{aligned} & -11.6511^{* * *} \\ & (0.0031) \end{aligned}$ |
| Blood test last year | - | - | - | $\begin{aligned} & -8.09870^{* *} \\ & (0.0030) \end{aligned}$ | $\begin{aligned} & -8.09865^{* *} \\ & (0.0030) \end{aligned}$ |
| Man | $\begin{aligned} & 13.3946 * * * \\ & (0.2152) \end{aligned}$ | $\begin{aligned} & 12.6625 * * * \\ & (0.2139) \end{aligned}$ | $\begin{aligned} & 12.4184^{* * *} \\ & (0.2215) \end{aligned}$ | $\begin{aligned} & 13.4264^{* * *} \\ & (0.2325) \end{aligned}$ | $\begin{aligned} & 13.1261^{* * *} \\ & (0.2322) \end{aligned}$ |
| Age | $\begin{aligned} & 7.8840^{* * *} \\ & (0.1381) \end{aligned}$ | $\begin{aligned} & 7.7335^{* * *} \\ & (0.1354) \end{aligned}$ | $\begin{aligned} & 7.6209 * * * \\ & (0.1505) \end{aligned}$ | $\begin{aligned} & 8.5209 * * * \\ & (0.1775) \end{aligned}$ | $\begin{aligned} & 8.2209 * * * \\ & (0.1662) \end{aligned}$ |
| Age $\wedge 2$ | $\begin{aligned} & -0.0495^{* * *} \\ & (0.0010) \end{aligned}$ | $\begin{aligned} & -0.0480^{* * *} \\ & (0.0010) \end{aligned}$ | $\begin{aligned} & -0.0473^{* * *} \\ & (0.0011) \end{aligned}$ | $\begin{aligned} & -0.0494^{* * *} \\ & (0.0022) \end{aligned}$ | $\begin{aligned} & -0.0491^{* * *} \\ & (0.0022) \end{aligned}$ |
| Austria | $\begin{aligned} & 8.2871^{* * *} \\ & (0.4195) \end{aligned}$ | $\begin{aligned} & 8.3152^{* * *} \\ & (0.4093) \end{aligned}$ | $\begin{aligned} & 8.3762^{* * *} \\ & (0.4138) \end{aligned}$ | $\begin{aligned} & 7.4952^{* * *} \\ & (0.4246) \end{aligned}$ | $\begin{aligned} & 7.1922 * * * \\ & (0.1216) \end{aligned}$ |
| Germany | $\begin{aligned} & 0.9488^{* *} \\ & (0.4315) \end{aligned}$ | $\begin{aligned} & 2.0698^{* * *} \\ & (0.4223) \end{aligned}$ | $\begin{aligned} & 2.1708^{* * *} \\ & (0.4249) \end{aligned}$ | $\begin{aligned} & 2.2906 * * * \\ & (0.4139) \end{aligned}$ | $\begin{aligned} & 2.2906 * * * \\ & (0.1139) \end{aligned}$ |
| Sweden | $\begin{aligned} & 10.0911^{* * *} \\ & (0.4284) \end{aligned}$ | $\begin{aligned} & 8.1773^{* * *} \\ & (0.4218) \end{aligned}$ | $\begin{aligned} & 8.3323^{* * *} \\ & (0.4486) \end{aligned}$ | $\begin{aligned} & 7.4424 * * * \\ & (0.4765) \end{aligned}$ | $\begin{aligned} & 7.1121^{* * *} \\ & (0.1662) \end{aligned}$ |
| Spain | $\begin{aligned} & 14.2872 * * * \\ & (0.4305) \end{aligned}$ | $\begin{aligned} & 15.2159 * * * \\ & (0.4203) \end{aligned}$ | $\begin{aligned} & 15.0797 * * * \\ & (0.4282) \end{aligned}$ | $\begin{aligned} & 16.0999 * * * \\ & (0.4162) \end{aligned}$ | $\begin{aligned} & 16.0999 * * * \\ & (0.1162) \end{aligned}$ |
| Italy | $\begin{aligned} & 8.2729 * * * \\ & (0.4241) \end{aligned}$ | $\begin{aligned} & 9.3606 * * * \\ & (0.4138) \end{aligned}$ | $\begin{aligned} & 9.3638^{* * *} \\ & (0.4176) \end{aligned}$ | $\begin{aligned} & 9.4546 * * * \\ & (0.4095) \end{aligned}$ | $\begin{aligned} & 9.1216^{* * *} \\ & (0.1092) \end{aligned}$ |
| France | $\begin{aligned} & -1.0237 * * \\ & (0.4081) \end{aligned}$ | $\begin{aligned} & -0.3794 \\ & (0.3981) \end{aligned}$ | $\begin{aligned} & -0.4362 \\ & (0.4006) \end{aligned}$ | $\begin{aligned} & -0.4570 \\ & (0.4126) \end{aligned}$ | $\begin{aligned} & -0.4265 \\ & (0.1126) \end{aligned}$ |
| Denmark | $\begin{aligned} & 18.1013 * * * \\ & (0.4402) \end{aligned}$ | $\begin{aligned} & 16.1364 * * * \\ & (0.4333) \end{aligned}$ | $\begin{aligned} & 16.4724^{* * *} \\ & (0.4563) \end{aligned}$ | $\begin{aligned} & 15.4967 * * * \\ & (0.4580) \end{aligned}$ | $\begin{aligned} & 14.1966 * * * \\ & (0.1280) \end{aligned}$ |
| Switzerland | $\begin{aligned} & 4.8486^{* * *} \\ & (0.4574) \end{aligned}$ | $\begin{aligned} & 3.7264^{* * *} \\ & (0.4472) \end{aligned}$ | $\begin{aligned} & 3.7549 * * * \\ & (0.4498) \end{aligned}$ | $\begin{aligned} & 4.0549^{* * *} \\ & (0.4599) \end{aligned}$ | $\begin{aligned} & 4.0219 * * * \\ & (0.1299) \end{aligned}$ |
| Wave 1-2 | $\begin{aligned} & -1.9374 * * * \\ & (0.2873) \end{aligned}$ | $\begin{aligned} & -1.9344^{* * *} \\ & (0.2995) \end{aligned}$ | $\begin{aligned} & -2.0010^{* * *} \\ & (0.3004) \end{aligned}$ | $\begin{aligned} & -2.3420^{* * *} \\ & (0.3713) \end{aligned}$ | $\begin{aligned} & -2.3120^{* * *} \\ & (0.3613) \end{aligned}$ |
| Wave 2-4 | $\begin{aligned} & -30.9290^{* * *} \\ & (0.2878) \end{aligned}$ | $\begin{aligned} & -30.2121^{* * *} \\ & (0.2979) \end{aligned}$ | $\begin{aligned} & -30.2523^{* * *} \\ & (0.2982) \end{aligned}$ | $\begin{aligned} & -32.6520^{* * *} \\ & (0.3089) \end{aligned}$ | $\begin{aligned} & -31.6120^{* * *} \\ & (0.3089) \end{aligned}$ |
| Wave 4-5 | $\begin{aligned} & 3.1023^{* * *} \\ & (0.3235) \end{aligned}$ | $\begin{aligned} & 3.0839 * * * \\ & (0.3301) \end{aligned}$ | $\begin{aligned} & 3.0392 * * * \\ & (0.3303) \end{aligned}$ | $\begin{aligned} & 3.1402 * * * \\ & (0.4204) \end{aligned}$ | $\begin{aligned} & 3.1102 * * * \\ & (0.1201) \end{aligned}$ |
| No education | $\begin{aligned} & 6.3875^{* * *} \\ & (0.5754) \end{aligned}$ | $\begin{aligned} & 9.5475 * * * \\ & (0.5662) \end{aligned}$ | $\begin{aligned} & 10.0616^{* * *} \\ & (0.5731) \end{aligned}$ | $\begin{aligned} & 11.1721^{* * *} \\ & (0.6042) \end{aligned}$ | $\begin{aligned} & 11.1621^{* * *} \\ & (0.6512) \end{aligned}$ |
| Primary education | $\begin{aligned} & 8.3876^{* * *} \\ & (0.3472) \end{aligned}$ | $\begin{aligned} & 10.9498^{* * *} \\ & (0.3432) \end{aligned}$ | $\begin{aligned} & 11.4072 * * * \\ & (0.3523) \end{aligned}$ | $\begin{aligned} & 12.5561^{* * *} \\ & (0.3624) \end{aligned}$ | $\begin{aligned} & 12.2261^{* * *} \\ & (0.3621) \end{aligned}$ |
| Lower secondary education | $\begin{aligned} & 3.3884^{* * *} \\ & (0.3547) \end{aligned}$ | $\begin{aligned} & 4.9869 * * * \\ & (0.3475) \end{aligned}$ | $\begin{aligned} & 5.3111^{* * *} \\ & (0.3527) \end{aligned}$ | $\begin{aligned} & 5.4222^{* * *} \\ & (0.4020) \end{aligned}$ | $\begin{aligned} & 5.1272 * * * \\ & (0.1020) \end{aligned}$ |
| Upper secondary education | $\begin{aligned} & 7.4730^{* * *} \\ & (0.2889) \end{aligned}$ | $\begin{aligned} & 8.6269 * * * \\ & (0.2827) \end{aligned}$ | $\begin{aligned} & 8.8285^{* * *} \\ & (0.2856) \end{aligned}$ | $\begin{aligned} & 7.6266^{* * *} \\ & (0.2970) \end{aligned}$ | $\begin{aligned} & 6.6266 * * * \\ & (0.2965) \end{aligned}$ |
| Post-secondary non-tertiary | $\begin{aligned} & 10.9202^{* * *} \\ & (0.6141) \end{aligned}$ | $\begin{aligned} & 11.3085 * * * \\ & (0.5975) \end{aligned}$ | $\begin{aligned} & 11.4665 * * * \\ & (0.5977) \end{aligned}$ | $\begin{aligned} & 12.4666 * * * \\ & (0.6045) \end{aligned}$ | $\begin{aligned} & 12.1666 * * * \\ & (0.6512) \end{aligned}$ |
| Married/cohabiting | $\begin{aligned} & 1.2261^{* * *} \\ & (0.3484) \end{aligned}$ | $\begin{aligned} & 0.6770^{* *} \\ & (0.3400) \end{aligned}$ | $\begin{aligned} & -1.0799 * * \\ & (0.4435) \end{aligned}$ | $\begin{aligned} & -1.2003^{* *} \\ & (0.5046) \end{aligned}$ | $\begin{aligned} & -1.2003^{* *} \\ & (0.2016) \end{aligned}$ |
| Single | $\begin{aligned} & -0.2741 \\ & (0.5362) \end{aligned}$ | $\begin{aligned} & -0.4163 \\ & (0.5221) \end{aligned}$ | $\begin{aligned} & -0.7104 \\ & (0.5236) \end{aligned}$ | $\begin{aligned} & -0.7704 \\ & (0.5246) \end{aligned}$ | $\begin{aligned} & -0.7651 \\ & (0.2216) \end{aligned}$ |
| Separated/Divorced | $\begin{aligned} & \text { 2.2170*** } \\ & (0.4726) \end{aligned}$ | $\begin{aligned} & 2.0405^{* * *} \\ & (0.4601) \end{aligned}$ | $\begin{aligned} & 1.7195^{* * *} \\ & (0.4628) \end{aligned}$ | $\begin{aligned} & 2.0206 * * * \\ & (0.456) \end{aligned}$ | $\begin{aligned} & 2.0211^{* * *} \\ & (0.126) \end{aligned}$ |
| Days in hospital (surgery/medical tests) |  | $\begin{aligned} & -0.1780^{* * *} \\ & (0.0138) \end{aligned}$ | $\begin{aligned} & -0.1797 * * * \\ & (0.0138) \end{aligned}$ | $\begin{aligned} & -0.2107^{* * *} \\ & (0.0146) \end{aligned}$ | $\begin{aligned} & -0.2106 * * * \\ & (0.0116) \end{aligned}$ |
| Days in hospital (mental health) |  | $\begin{aligned} & -0.1118^{* * *} \\ & (0.0546) \end{aligned}$ | $\begin{aligned} & -0.1120^{* * *} \\ & (0.0541) \end{aligned}$ | $\begin{aligned} & -0.1220^{* * *} \\ & (0.0142) \end{aligned}$ | $\begin{aligned} & -0.1320^{* * *} \\ & (0.0112) \end{aligned}$ |
| Days in other institutions |  | $\begin{aligned} & 0.0645^{* * *} \\ & (0.0215) \end{aligned}$ | $\begin{aligned} & 0.0642^{* * *} \\ & (0.0209) \end{aligned}$ | $\begin{aligned} & 0.0682 * * * \\ & (0.0222) \end{aligned}$ | $\begin{aligned} & 0.0690^{* * *} \\ & (0.0222) \end{aligned}$ |
| Vistis to GP |  | $\begin{aligned} & -0.0903^{* * *} \\ & (0.0223) \end{aligned}$ | $\begin{aligned} & -0.0864^{* * *} \\ & (0.0223) \end{aligned}$ | $\begin{aligned} & -0.0891^{* * *} \\ & (0.0244) \end{aligned}$ | $\begin{aligned} & -0.0875 * * * \\ & (0.0211) \end{aligned}$ |
| Employed |  |  | $\begin{aligned} & 0.7910^{* *} \\ & (0.3805) \end{aligned}$ | $\begin{aligned} & 0.8020^{* *} \\ & (0.4006) \end{aligned}$ | $\begin{aligned} & 0.8033^{* *} \\ & (0.1006) \end{aligned}$ |
| Unemployed |  |  | $\begin{aligned} & 1.1400^{*} \\ & (0.6498) \end{aligned}$ | $\begin{aligned} & 1.2400^{*} \\ & (0.6506) \end{aligned}$ | $\begin{aligned} & 1.2100^{*} \\ & (0.6206) \end{aligned}$ |
| Retired |  |  | $\begin{aligned} & 1.3713^{* * *} \\ & (0.3353) \end{aligned}$ | $\begin{aligned} & 1.4024^{* * *} \\ & (0.3464) \end{aligned}$ | $\begin{aligned} & 1.1221^{* * *} \\ & (0.3161) \end{aligned}$ |
| Lives alone |  |  | $\begin{aligned} & -2.4425 * * * \\ & (0.3972) \end{aligned}$ | $\begin{aligned} & -2.4711^{* * *} \\ & (0.4002) \end{aligned}$ | $\begin{aligned} & -2.4611^{* * *} \\ & (0.1002) \end{aligned}$ |
| Adjusted income (1,000 PPP) |  |  | $\begin{aligned} & 66.417 * * \\ & (30.565) \end{aligned}$ | $\begin{aligned} & 66.420^{* *} \\ & (32.666) \end{aligned}$ | $\begin{aligned} & 66.390^{* *} \\ & (32.676) \end{aligned}$ |
| Adjusted wealth (1,000 PPP) |  |  | $\begin{aligned} & 1.1664 * * * \\ & (0.3838) \end{aligned}$ | $\begin{aligned} & 1.2651 * * * \\ & (0.3646) \end{aligned}$ | $\begin{aligned} & 1.2621^{* * *} \\ & (0.3716) \end{aligned}$ |
| Rural área, small village |  |  | $\begin{aligned} & -0.8729 * * * \\ & (0.2135) \end{aligned}$ | $\begin{aligned} & -0.9020^{* * *} \\ & (0.2246) \end{aligned}$ | $\begin{aligned} & -0.9110^{* * *} \\ & (0.116) \end{aligned}$ |
| Constant | $\begin{aligned} & -29.6850^{* * *} \\ & (4.006) \\ & \hline \end{aligned}$ | $\begin{aligned} & -29.8698^{* * *} \\ & (4.172) \\ & \hline \end{aligned}$ | $\begin{aligned} & -29.6701^{* * *} \\ & (5.532) \\ & \hline \end{aligned}$ | $\begin{aligned} & -12.6002^{* * *} \\ & (5.781) \\ & \hline \end{aligned}$ | $\begin{aligned} & -12.65102^{* * *} \\ & (5.681) \\ & \hline \end{aligned}$ |
| N | 69,990 | 69,990 | 69,990 | 33,269 | 33,269 |
| $\mathrm{R}^{2}$ | 0.3466 | 0.3818 | 0.3825 | 0.3712 | 0.3612 |


$\frac{\mathrm{p}}{\text { Omitted variables: women, tertiary education, widow, relation with economic activity: houseworking and other situations, Belgium, waves } 5-6 \text {. Bootstrap with }}$ 1,000 replications. Robust standard errors between parenthesis. * statistically significant at $10 \%$; ** statistically significant at $5 \%$; *** statistically significant at $1 \%$. Income and wealth are expressed in 1,000 units of Purchase Power of Parity and adjusted by household sized (dividing wealth and income by the square root of the number of household members).
${ }^{\text {a }}$ Recorded father's age at decease if father has died at the time of the survey. Predicted father's age at decease if father is alive at the time of the survey.
Recorded mother's age at decease if mother has died at the time of the survey. Predicted mother's age at decease if mother is alive at the time of the survey.

Table D2. Effect of the instruments over the dependent variables. Reduced form equations.

|  | Smokes now | Has never smoked | Drinks every day or 5/6 times per week | Has not drunk during last 3 months or less than once per month |
| :---: | :---: | :---: | :---: | :---: |
| Father's age at decease | $\begin{aligned} & \hline-0.0006 \\ & (0.0006) \end{aligned}$ | $\begin{aligned} & \hline 0.0039 \\ & (0.0027) \end{aligned}$ | $\begin{aligned} & \hline 0.0005 \\ & (0.0007) \end{aligned}$ | $\begin{aligned} & \hline-0.0006 \\ & (0.0007) \end{aligned}$ |
| Mother's age at decease | $\begin{aligned} & 0.0003 \\ & (0.0006) \end{aligned}$ | $\begin{aligned} & -0.0040 \\ & (0.0037) \end{aligned}$ | -0.0000 $(0.0007)$ | $\begin{aligned} & -0.0002 \\ & (0.0007) \end{aligned}$ |
| Father alive | $\begin{aligned} & -0.0230^{* *} \\ & (0.0105) \end{aligned}$ | $\begin{aligned} & -0.0105 \\ & (0.0121) \end{aligned}$ | $\begin{aligned} & -0.0167 \\ & (0.0122) \end{aligned}$ | $\begin{aligned} & -0.0023 \\ & (0.0116) \end{aligned}$ |
| Mother alive | $\begin{aligned} & -0.0229 \\ & (0.0182) \end{aligned}$ | $\begin{aligned} & 0.1216 \\ & (0.0741) \end{aligned}$ | $\begin{aligned} & 0.0261^{*} \\ & (0.0141) \end{aligned}$ | $\begin{aligned} & -0.0087 \\ & (0.0134) \end{aligned}$ |
| N | 69,990 | 69,990 | 69,990 | 69,990 |
| r2 | 0.0754 | 0.2568 | 0.1241 | 0.2198 |
| F | 215.858 | 914.221 | 374.888 | 745.457 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | Sedentary lifestyle | Normal weight | Obese or overweight | Ten-word listing (first recall) |
| Father's age at decease | $\begin{aligned} & 0.0007 \\ & (0.0005) \end{aligned}$ | $\begin{aligned} & -0.0008 \\ & (0.0008) \end{aligned}$ | $\begin{aligned} & 0.0010 \\ & (0.0006) \end{aligned}$ | $\begin{aligned} & -0.0007 \\ & (0.0026) \end{aligned}$ |
| Mother's age at decease | $\begin{aligned} & -0.0011^{* *} \\ & (0.0005) \end{aligned}$ | $\begin{aligned} & 0.0008 \\ & (0.0008) \end{aligned}$ | $\begin{aligned} & -0.0009 \\ & (0.0006) \end{aligned}$ | $\begin{aligned} & 0.0038 \\ & (0.0026) \end{aligned}$ |
| Father alive | -0.0079 | 0.0026 | -0.0051 | 0.0503 |
|  | (0.0082) | (0.0138) | (0.0110) | (0.0443) |
| Mother alive | 0.0034 | 0.0095 | -0.0166 | -0.0561 |
|  | (0.0095) | (0.0161) | (0.0128) | (0.0515) |
| N | 69,990 | 69,990 | 69,990 | 69,990 |
| r2 | 0.2198 | 0.0839 | 0.0838 | 0.3328 |
| F | 745.374 | 242.317 | 242.057 | 1.319.381 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | Ten word listing (second recall) | Maximum grip strength measure |  |  |
| Father's age at decease | -0.0059* | 0.0053 |  |  |
|  | (0.0031) | (0.0120) |  |  |
| Mother's age at decease | $0.0086$ | $0.0030$ |  |  |
| Father alive | 0.1507 | 0.2193 |  |  |
|  | (0.0927) | (0.2076) |  |  |
| Mother alive | -0.1097* | -0.0897 |  |  |
|  | (0.0613) | (0.2414) |  |  |
| N | 69,990 | 69,990 |  |  |
| r2 | 0.3030 | 0.6691 |  |  |
| F | 1,150.028 | 5,027.791 |  |  |
| p | 0.0000 | 0.0000 |  |  |
|  | Bank accounts | Stocks or shares (listed or unlisted in stockmarket) | Government or corporate bonds | Mutual funds |
| Father's age at decease | -0.0003 | -0.0009** | -0.0002 | -0.0004 |
|  | (0.0004) | (0.0005) | (0.0003) | (0.0005) |
| Mother's age at decease | 0.0006 | 0.0001 | 0.0004 | 0.0007 |
|  | (0.0004) | (0.0005) | (0.0003) | (0.0005) |
| Father alive | -0.0056 | 0.0023 | 0.0032 | 0.0120 |
|  | (0.0068) | (0.0077) | (0.0057) | (0.0085) |
| Mother alive | 0.0100 | -0.0112 | -0.0087 | -0.0149 |
|  | (0.0079) | (0.0090) | (0.0067) | (0.0099) |
| N | 69,990 | 69,990 | 69,990 | 69,990 |
| r2 | 0.1812 | 0.1150 | 0.0541 | 0.1008 |
| F | 585.315 | 343.805 | 151.300 | 296.596 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | Mutual funds: mostly in stocks | Mutual funds: half in stocks, half in bonds | Mutual funds: mostly in bonds | Individual retirement accounts |
| Father's age at decease | -0.0013 | -0.0038 | 0.0025 | 0.0005 |
|  | (0.0026) | (0.0028) | (0.0033) | (0.0005) |
| Mother's age at decease | 0.0004 | 0.0038 | -0.0034 | -0.0002 |
|  | (0.0026) | (0.0028) | (0.0022) | (0.0005) |
| Father alive | 0.0186 | -0.0027 | -0.0160 | $0.0011$ |
|  | (0.0388) | (0.0421) | (0.0345) | (0.0080) |
| Mother alive | 0.0334 | 0.0258 | -0.0591 | -0.0159* |
|  | (0.0448) | (0.0487) | (0.0399) | (0.0094) |
| N | 69,990 | 69,990 | 69,990 | 69,990 |
| r2 | 0.1698 | 0.1026 | 0.0954 | 0.1379 |
| F | 58.897 | 32.929 | 30.375 | 423.097 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | Life insurance | Mortgage | Other debts | Total wealth |
| Father's age at decease | -0.0002 | 0.0003 | 0.0011 | 0.0004 |


| Mother's age at decease | (0.0005) | (0.0005) | (0.0007) | (0.0004) |
| :---: | :---: | :---: | :---: | :---: |
|  | 0.0004 | -0.0001 | -0.0012 | 0.0000 |
|  | (0.0005) | (0.0005) | (0.0008) | (0.0004) |
| Father alive | -0.0009 | 0.0153 | -0.0143* | -0.0052 |
|  | (0.0091) | (0.0094) | (0.0086) | (0.0075) |
| Mother alive | -0.0082 | -0.0161 | 0.0136 | -0.0199 |
|  | (0.0106) | (0.0109) | (0.0100) | (0.0108) |
| N | 69,990 | 69,990 | 69,990 | 69,990 |
| r2 | 0.1182 | 0.1711 | 0.0970 | 0.1306 |
| F | 354.477 | 546.123 | 284.024 | 397.560 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 |
|  | Take substantial risk expecting to earn substantial returns | Take above average financial risks expecting to earn above average returns | Take average financial risks expecting to earn average returns | Not willing to take any financial risks |
| Father's age at decease | 0.0000 | -0.0003 | 0.0004 | 0.0000 |
|  | (0.0001) | (0.0002) | (0.0006) | (0.0001) |
| Mother's age at decease | -0.0001 | 0.0003 | -0.0002 | -0.0001 |
|  | (0.0001) | (0.0002) | (0.0006) | (0.0001) |
| Father alive | -0.0017 | 0.0047 | 0.0091 | -0.0017 |
|  | (0.0024) | (0.0042) | (0.0096) | (0.0024) |
| Mother alive | 0.0019 | -0.0007 | -0.0175 | 0.0019 |
|  | (0.0028) | (0.0049) | (0.0111) | (0.0028) |
| N | 69,990 | 69,990 | 69,990 | 69,990 |
| r2 | 0.0126 | 0.0449 | 0.1680 | 0.0126 |
| F | 33.628 | 124.255 | 534.321 | 33.628 |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 |

All regressions include gender, age, age squared, level of education (no education, primary education, lower secondary education, higher secondary education, post-secondary non-tertiary education and tertiary education (omitted)), marital status (married/cohabiting, separated/divorced, single and widow (omitted)), number of days stayed at hospital during last year (surgery or medical tests), number of days stayed at hospital during last year (mental health), number of days stayed at other institutions during last year, number of visits to general practitioner during last year, relation with economic activity (employed, unemployed, retired, houseworking and other situations (omitted), living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in $1,000 \mathrm{PPP}$ ), country fixed effects and wave fixed effects.

Table D3. Comparison of cognitive and economic related variables between focal respondents and non-focal respondents

|  | Focal response $=0$ | Focal response $=50$ | Focal response $=100$ | No focal response |
| :---: | :---: | :---: | :---: | :---: |
| Deceased subsample Cognitive skills |  |  |  |  |
| Ten word listing (first recall) | 2.87 | 3.56 | 3.70 | 3.81 |
|  | (2.01) | (2.04) | (2.08) | (1.91) |
| Ten word listing (second recall) | 1.60 | 2.00 | 1.98 | 2.33 |
|  | (1.77) | (1.89) | (1.91) | (1.94 |
| Education (ISCED 1997 levels) |  |  |  |  |
| None | 15.18 | 14.59 | 12.79 | 11.24 |
| Primary education | 40.44 | 42.60 | 35.89 | 40.01 |
| Lower secondary | 14.69 | 12.43 | 6.44 | 14.77 |
| Upper secondary | 19.52 | 19.84 | 30.38 | 23.71 |
| Post-secondary non tertiary | 1.94 | 1.00 | 3.22 | 1.34 |
| Tertiary education | 8.22 | 9.54 | 11.28 | 8.94 |
| Adjusted income (1,000 PPP) ${ }^{\text {a }}$ | 14.03 | 15.73 | 20.40 | 15.09 |
|  | (15.93) | (19.51) | (26.24 | (22.81) |
| Adjusted wealth (1,000 PPP) ${ }^{\text {a }}$ | 112.87 | 138.14 | 128.63 | 128.72 |
|  | (161.91) | (172.22) | (183.44 | (206.44) |
| N | 350 | 385 | 139 | 1,166 |
| Survivors subsample |  |  |  |  |
| Cognitive skills |  |  |  |  |
| Ten word listing (first recall) | 4.01 | 4.99 | 5.30 | 5.05 |
|  | (1.88) | (1.80) | (1.73) | (1.81) |
| Ten word listing (second recall) | 2.53 | 3.54 | 3.91 | 3.64 |
|  | (1.92) | (2.05) | (2.05) | (2.06) |
| Education (ISCED 1997 levels) ${ }^{\text {l }}$ ( ${ }^{\text {a }}$ |  |  |  |  |
| None | 10.09 | 6.30 | 4.35 | 6.30 |
| Primary education | 32.28 | 24.27 | 21.68 | 23.96 |
| Lower secondary | 16.43 | 15.34 | 16.34 | 16.11 |
| Upper secondary | 28.16 | 32.69 | 34.76 | 30.86 |
| Post-secondary non tertiary | 1.32 | 2.40 | 3.10 | 2.33 |
| Tertiary education | 11.72 | 19.00 | 19.78 | 20.44 |
| Adjusted income (1,000 PPP) ${ }^{\text {a }}$ | 15.73 | 20.54 | 20.43 | 20.24 |
|  | (16.68) | (26.91) | (22.32) | (26.72) |
| Adjusted wealth (1,000 PPP) ${ }^{\text {a }}$ | 126.49 | 163.95 | 178.94 | 170.41 |
|  | (224.37) | (246.44) | (309.62) | (301.22) |
| N | 2,420 | 14,152 | 11,561 | 39,727 |

${ }^{a}$ Income and wealth are expressed in 1,000 units of Purchase Power of Parity and adjusted by household sized (dividing wealth and income by the square root of the number of household members).Standard deviation between parenthesis.

For the deceased subsample and compared to non-focal respondents we show that: (i) focal respondents at zero attain lower scores for both ten-word listing recalls, have a higher percentage without primary education and lower wealth, (ii) focal respondents at 100 attain a lower score for the second ten-word listing recall, have a higher percentage with post-secondary education, higher income and higher wealth. For the survivors' subsample, and compared to non-focal respondents: (i) focal respondents at 100 achieve higher scores for both ten-word listing recalls and show similar levels of income and wealth; (ii) the opposite is true for focal respondents at 0 . For both subsamples, ten-word listing scores, income and wealth of focal-respondents at 50 resemble to those of non-focal respondents. However, the distribution of educational levels is more akin to that of focal respondents at 0 for the deceased subsample and more similar to that of non-focal respondents for the survivors' subsample.

Table D4. Subsample of respondents who have answered SHARELIFE. Replication of regressions shown on Tables B.13 to B.17.
Regressions after instrumenting the "biased survival" indicator" (difference between subjective survival probability and objective survival probability) introducing as instruments blood pressure checked last year and blood test done last year.
Reported coefficients correspond to the variable "Biased survival expectations" after IV regression.

|  | LO > 0 |  |  | LO < 0 |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Dependent variable | M1 | M2 | M3 | M1 | M2 | M3 |
| Smokes now | $\begin{array}{r} \hline-0.6090^{* *} \\ (0.2317) \end{array}$ | $\begin{array}{r} -0.5995^{* *} \\ (0.2001) \end{array}$ | $\begin{array}{r} \hline-0.5245^{* *} \\ (0.1943) \end{array}$ | $\begin{array}{r} \hline 0.8714^{* * *} \\ (0.2102) \end{array}$ | $\begin{array}{r} \hline 0.8585^{* * *} \\ (0.2027) \end{array}$ | $\begin{array}{r} 0.8515^{* * *} \\ (0.2024) \end{array}$ |
| Has never smoked | $\begin{array}{r} 1.1568^{* *} \\ (0.4634) \end{array}$ | $\begin{array}{r} 1.0988^{* *} \\ (0.4635) \end{array}$ | $\begin{array}{r} 1.0811^{* *} \\ (0.4512) \end{array}$ | $\begin{array}{r} -1.6464^{* * *} \\ (0.4715) \end{array}$ | $\begin{array}{r} -1.4321^{* * *} \\ (0.4792) \end{array}$ | $\begin{array}{r} -1.3744^{* * *} \\ (0.4770) \end{array}$ |
| Drinks every day or 5/6 times per week | $\begin{array}{r} -0.8751^{* * *} \\ (0.2307) \end{array}$ | $\begin{array}{r} -0.8312 * * * \\ (0.2126) \end{array}$ | $\begin{array}{r} -0.7868^{* * *} \\ (0.2098) \end{array}$ | $\begin{array}{r} 1.0739 * * * \\ (0.3094) \end{array}$ | $\begin{array}{r} 0.9839 * * * \\ (0.3057) \end{array}$ | $\begin{array}{r} 0.9451^{* * *} \\ (0.3141) \end{array}$ |
| Has not drunk during last 3 months or less than once per month | $\begin{gathered} 1.3329 * * \\ (0.5750) \end{gathered}$ | $\begin{array}{r} 1.3016^{* *} \\ (0.5707) \end{array}$ | $\begin{gathered} 1.2598^{* *} \\ (0.5712) \end{gathered}$ | $\begin{array}{r} -0.5870^{* *} \\ (0.1964) \end{array}$ | $\begin{array}{r} -0.4688^{* *} \\ (0.1839) \end{array}$ | $\begin{array}{r} -0.4648^{* *} \\ (0.1445) \end{array}$ |
| Sedentary lifestyle | $\begin{array}{r} -0.9664^{* *} \\ (0.4232) \end{array}$ | $\begin{array}{r} -0.9621^{* *} \\ (0.4241) \end{array}$ | $\begin{array}{r} -0.9588^{* *} \\ (0.4144) \end{array}$ | $\begin{array}{r} 1.0116^{* *} \\ (0.3551) \end{array}$ | $\begin{array}{r} 0.8798^{* *} \\ (0.3401) \end{array}$ | $\begin{array}{r} 0.8776 * * \\ (0.3060) \end{array}$ |
| Normal weight | $\begin{array}{r} 0.9436^{* * *} \\ (0.2305) \end{array}$ | $\begin{array}{r} 0.7906 * * * \\ (0.2569) \end{array}$ | $\begin{array}{r} 0.7646 * * * \\ (0.2342) \end{array}$ | $\begin{array}{r} -0.9584^{* *} \\ (0.3032) \end{array}$ | $\begin{array}{r} -0.7055 * * \\ (0.2773) \end{array}$ | $\begin{array}{r} -0.6366^{* *} \\ (0.2668) \end{array}$ |
| Obese | $\begin{array}{r} -1.1377 * * \\ (0.4038) \end{array}$ | $\begin{gathered} -1.1200^{*} \\ (0.3644) \end{gathered}$ | $\begin{array}{r} -1.0790^{* *} \\ (0.3634) \end{array}$ | $\begin{array}{r} 1.9615^{* * *} \\ (0.5693) \end{array}$ | $\begin{array}{r} 1.6149 * * \\ (0.5746) \end{array}$ | $\begin{array}{r} 1.2987 * * \\ (0.5704) \end{array}$ |
| Ten word listing first recall | $\begin{array}{r} 3.8851^{* * *} \\ (0.7363) \end{array}$ | $\begin{array}{r} 3.2821^{* * *} \\ (0.7277) \end{array}$ | $\begin{array}{r} 2.3868^{* * *} \\ (0.7095) \end{array}$ | $\begin{array}{r} -0.6722^{* * *} \\ (0.1599) \end{array}$ | $\begin{array}{r} -0.6570^{* * *} \\ (0.1968) \end{array}$ | $\begin{array}{r} -0.6392^{* * *} \\ (0.1870) \end{array}$ |
| Ten word listing second recall | $\begin{array}{r} 3.1237 * * * \\ (0.7385) \end{array}$ | $\begin{array}{r} 2.5025 * * * \\ (0.7332) \end{array}$ | $\begin{array}{r} 2.5128 * * * \\ (0.7314) \end{array}$ | $\begin{array}{r} -2.1674^{* *} \\ (0.8193) \end{array}$ | $\begin{array}{r} -1.6763^{* *} \\ (0.7618) \end{array}$ | $\begin{array}{r} -1.7006^{* *} \\ (0.6750) \end{array}$ |
| Maximum grip strength measure | $\begin{array}{r} 7.4799 * * * \\ (2.3120) \end{array}$ | $\begin{array}{r} 5.8987 * * \\ (2.9973) \end{array}$ | $\begin{array}{r} 5.8917 * * \\ (2.4030) \end{array}$ | $\begin{array}{r} -6.7695^{* *} \\ (2.9708) \end{array}$ | $\begin{array}{r} -6.6219 * * \\ (2.8563) \end{array}$ | $\begin{array}{r} -6.2504^{* *} \\ (2.4591) \end{array}$ |
| Social contacts | $\begin{array}{r} 0.2854^{* * *} \\ (0.0791) \end{array}$ | $\begin{array}{r} 0.2829 * * * \\ (0.0701) \end{array}$ | $\begin{array}{r} 0.2540^{* * *} \\ (0.0615) \end{array}$ | $\begin{array}{r} 0.0581^{* *} \\ (0.0166) \end{array}$ | $\begin{array}{r} -0.0563^{* * *} \\ (0.0121) \end{array}$ | $\begin{array}{r} -0.0560^{* * *} \\ (0.0125) \end{array}$ |
| Bank accounts | $\begin{array}{r} -0.7356^{* * *} \\ (0.1944) \end{array}$ | $\begin{array}{r} -0.7247 * * * \\ (0.2232) \end{array}$ | $\begin{array}{r} -0.6930^{* * *} \\ (0.2162) \end{array}$ | $\begin{array}{r} 0.9972 * * * \\ (0.1806) \end{array}$ | $\begin{array}{r} 0.9707 * * * \\ (0.2080) \end{array}$ | $\begin{array}{r} 0.9596^{* * *} \\ (0.2021) \end{array}$ |
| Government or corporate bonds | $\begin{gathered} 0.0955^{* *} \\ (0.0379) \end{gathered}$ | $\begin{array}{r} 0.0946 * * \\ (0.0393) \end{array}$ | $\begin{gathered} 0.0849^{* *} \\ (0.0326) \end{gathered}$ | $\begin{array}{r} -0.1676^{* *} \\ (0.0836) \end{array}$ | $\begin{array}{r} -0.1526^{* *} \\ (0.0779) \end{array}$ | $\begin{array}{r} -0.1495^{* *} \\ (0.0776) \end{array}$ |
| Stocks or shares listed or unlisted in stockmarket | $\begin{array}{r} -0.0982 * * \\ (0.0041) \end{array}$ | $\begin{array}{r} -0.0946 * * \\ (0.0037) \end{array}$ | $\begin{array}{r} -0.0926^{* *} \\ (0.0034) \end{array}$ | $\begin{array}{r} 0.1633^{* * *} \\ (0.0528) \end{array}$ | $\begin{array}{r} 0.1597 * * * \\ (0.0511) \end{array}$ | $\begin{gathered} 0.1562 * * \\ (0.0493) \end{gathered}$ |
| Mutual funds or managed investment accounts | $\begin{array}{r} 0.1761^{* * *} \\ (0.0505) \end{array}$ | $\begin{array}{r} 0.1689^{* * *} \\ (0.0491) \end{array}$ | $\begin{array}{r} 0.1574^{* * *} \\ (0.0488) \end{array}$ | $\begin{array}{r} -0.2836 * * * \\ (0.0888) \end{array}$ | $\begin{array}{r} -0.2772 * * * \\ (0.0866) \end{array}$ | $\begin{array}{r} -0.2586 * * * \\ ((0.0833) \end{array}$ |
| Mutual funds: mostly in stocks | $\begin{array}{r} -0.2624^{* *} \\ (0.1274) \end{array}$ | $\begin{array}{r} -0.2408^{* *} \\ (0.1087) \end{array}$ | $\begin{array}{r} -0.2254^{* *} \\ (0.1012) \end{array}$ | $\begin{array}{r} 0.2137 * * * \\ (0.0576) \end{array}$ | $\begin{array}{r} 0.1916 * * * \\ (0.0585) \end{array}$ | $\begin{array}{r} 0.1229 * * * \\ (0.0583) \end{array}$ |
| Mutual funds: half in stocks. half in bonds | $\begin{array}{r} -0.3769^{* * *} \\ (0.0997) \end{array}$ | $\begin{array}{r} -0.3371^{* * *} \\ (0.0958) \end{array}$ | $\begin{array}{r} -0.3600^{* * *} \\ (0.0890) \end{array}$ | $\begin{array}{r} 1.3707^{* * *} \\ (0.3323) \end{array}$ | $\begin{array}{r} 1.3524^{* * *} \\ (0.3775) \end{array}$ | $\begin{array}{r} 1.2609^{* * *} \\ (0.3554) \end{array}$ |
| Mutual funds: mostly in bonds | $\begin{array}{r} 0.5659 * * \\ (0.2085) \end{array}$ | $\begin{gathered} 0.5115^{* *} \\ (0.2209) \end{gathered}$ | $\begin{array}{r} 0.4626^{* *} \\ (0.1926) \end{array}$ | $\begin{array}{r} -0.5094^{* *} \\ (0.2347) \end{array}$ | $\begin{array}{r} -0.5033^{* *} \\ (0.2271) \end{array}$ | $\begin{array}{r} -0.5021^{* *} \\ (0.2131) \end{array}$ |
| Individual retirement accounts | $\begin{array}{r} 0.2341^{* * *} \\ (0.0423) \end{array}$ | $\begin{array}{r} 0.2244^{* * *} \\ (0.0476) \end{array}$ | $\begin{array}{r} 0.2102^{* * *} \\ (0.0482) \end{array}$ | $\begin{array}{r} -0.3356^{* * *} \\ (0.0994) \end{array}$ | $\begin{array}{r} -0.3641^{* * *} \\ (0.0955) \end{array}$ | $\begin{array}{r} -0.3200^{* * *} \\ (0.0921) \end{array}$ |
| Life insurance | $\begin{array}{r} -0.2734^{* * *} \\ (0.0729) \end{array}$ | $\begin{array}{r} -0.2341^{* * *} \\ (0.0714) \end{array}$ | $\begin{array}{r} -0.2146^{* * *} \\ (0.0649) \end{array}$ | $\begin{array}{r} 0.3747 * * \\ (0.1273) \end{array}$ | $\begin{gathered} 0.3634^{* *} \\ (0.1217) \end{gathered}$ | $\begin{array}{r} 0.3478 * * \\ (0.1224) \end{array}$ |
| Mortgage | $\begin{array}{r} 0.1650 * * * \\ (0.0464) \end{array}$ | $\begin{array}{r} 0.1627 * * * \\ (0.0459) \end{array}$ | $\begin{array}{r} 0.1584^{* * *} \\ (0.0454) \end{array}$ | $\begin{array}{r} -0.1493^{* * *} \\ (0.0632) \end{array}$ | $\begin{array}{r} -0.1434^{* *} \\ (0.0629) \end{array}$ | $\begin{gathered} -0.1296 * \\ (0.0570) \end{gathered}$ |
| Other debts | $\begin{gathered} 0.1611^{* *} \\ (0.0586) \end{gathered}$ | $\begin{array}{r} 0.1575 * * \\ (0.0521) \end{array}$ | $\begin{array}{r} 0.1446 * * \\ (0.0528) \end{array}$ | $\begin{array}{r} -0.1627 * * \\ (0.0772) \end{array}$ | $\begin{array}{r} -0.1610^{* *} \\ (0.0715) \end{array}$ | $\begin{array}{r} -0.1580^{* *} \\ (0.0728) \end{array}$ |
| Total wealth amount 1.000 PPP | $\begin{array}{r} 0.2362^{* * *} \\ (0.0942) \end{array}$ | $\begin{array}{r} 0.2158^{* *} \\ (0.0952) \end{array}$ | $\begin{array}{r} 0.2168^{* * *} \\ (0.0946) \end{array}$ | $\begin{array}{r} -0.7678^{* *} \\ (0.2287) \end{array}$ | $\begin{array}{r} -0.6672^{* *} \\ (0.2168) \end{array}$ | $\begin{array}{r} -0.6247 * * \\ (0.2128) \end{array}$ |
| Take substantial risk expecting to earn substantial returns | $\begin{array}{r} 0.0277 * * * \\ (0.0035) \end{array}$ | $\begin{array}{r} 0.0251^{* * *} \\ (0.0044) \end{array}$ | $\begin{array}{r} 0.0248 * * * \\ (0.0022) \end{array}$ | $\begin{array}{r} -0.0386^{* * *} \\ (0.0021) \end{array}$ | $\begin{array}{r} -0.0302 * * * \\ (0.0020) \end{array}$ | $\begin{array}{r} -0.0282 * * * \\ (0.0079) \end{array}$ |
| Take above average financial risks expecting to earn above average returns | $\begin{array}{r} 1.1838 * * * \\ (0.3607) \end{array}$ | $\begin{array}{r} 1.1639 * * * \\ (0.4112) \end{array}$ | $\begin{array}{r} 1.1229 * * * \\ (0.3955) \end{array}$ | $\begin{array}{r} -0.9926 * * \\ (0.3647) \end{array}$ | $\begin{array}{r} -0.9406 * * \\ (0.4118) \end{array}$ | $\begin{array}{r} -0.9136 * * \\ (0.3973) \end{array}$ |
| Take average financial risks expecting to earn average returns | $\begin{array}{r} -0.0768^{* * *} \\ (0.0211) \end{array}$ | $\begin{array}{r} -0.0683^{* * *} \\ (0.0208) \end{array}$ | $\begin{array}{r} -0.0671^{* * *} \\ (0.0205) \end{array}$ | $\begin{array}{r} 0.1086 * * * \\ (0.0065) \end{array}$ | $\begin{array}{r} 0.1066^{* * *} \\ (0.0083) \end{array}$ | $\begin{array}{r} 0.1020^{* * *} \\ (0.0017) \end{array}$ |
| Not willing to take any financial risks | $\begin{array}{r} -0.6589 * * * \\ (0.1420) \\ \hline \end{array}$ | $\begin{array}{r} -0.5587 * * * \\ (0.1662) \\ \hline \end{array}$ | $\begin{array}{r} -0.5441^{* * *} \\ (0.1739) \\ \hline \end{array}$ | $\begin{array}{r} 0.4109^{* * *} \\ (0.0493) \\ \hline \end{array}$ | $\begin{array}{r} 0.3874 * * * \\ (0.0786) \\ \hline \end{array}$ | $\begin{array}{r} 0.3568 * * * \\ (0.0865) \\ \hline \end{array}$ |
| Education, marital status | Yes | Yes | Yes | Yes | Yes | Yes |
| Days at hospital for surgery/tests, due to mental problems, days at other institutions, visits to GP <br> Relation with economic activity, lives alone, lives in rural area, income and wealth <br> Country fixed effects <br> Time fixed effects | No <br> No <br> Yes <br> Yes | Yes <br> No <br> Yes <br> Yes | Yes <br> Yes <br> Yes <br> Yes | No <br> No <br> Yes <br> Yes | Yes <br> No <br> Yes <br> Yes | Yes <br> Yes <br> Yes <br> Yes |

M1 includes optimistic indicator (after first-stage regression), gender, age, age squared, level of education (no education, primary education, lower secondary education, higher secondary education, post-secondary non-tertiary education and tertiary education (omitted)), marital status (married/cohabiting, separated/divorced, single and widow (omitted)), country fixed effects and wave fixed effects. M2 includes the same explanatory variables than M1, and additionally, number of days stayed at hospital during last year (surgery or medical tests), number of days stayed at hospital during last year (mental health), number of days stayed at other institutions during last year and number of visits to general practitioner during last year. M3 includes the same explanatory variables than M2, an additionally, relation with economic activity (employed, unemployed, retired, houseworking and other situations (omitted)), living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in 1,000 PPP). Bootstrap with 1,000 replications Robust standard errors between parenthesis. * statistically significant at 10\%; ** statistically significant at 5\%; *** statistically significant at $1 \%$.
Results of the Durbin test, Stock and Yogo test and Sargan test are not shown due to space constraints but are available upon request. Results indicate that OLS estimations are not consistent and confirm the validity of the instruments used.
$\mathrm{N}=33,269$; 21,313 for the sample with $\mathrm{LO}>0$ and 11,956 for the sample with $\mathrm{LO}<0$.

Table D5. Comparison between current and lagged instruments in first-stage regression for "biased survival"

|  | Using the whole sample of respondents | Sample of respondents who have a lagged observations |  |
| :---: | :---: | :---: | :---: |
|  |  | With current instruments | With lagged instruments |
| Mother's age at decease ${ }^{\text {a }}$ | $0.0012 * * *$ (0.0001) | 0.0013*** |  |
| Father's age at decease ${ }^{\text {b }}$ | $0.0017 * * *$ | 0.0019*** |  |
|  | (0.0001) | (0.0001) |  |
| Mother alive | $\begin{aligned} & 0.0114^{* *} \\ & (0.0050) \end{aligned}$ | $\begin{aligned} & 0.0127^{* *} \\ & (0.0056) \end{aligned}$ |  |
| Father alive | $0.0248 * * *$ | 0.0275*** |  |
| Mother's age at decease (lagged) ${ }^{\text {a }}$ |  |  | $\begin{aligned} & 0.0017 * * * \\ & (0.0001) \end{aligned}$ |
| Father's age at decease (lagged) ${ }^{\text {b }}$ |  |  | $\begin{aligned} & 0.0024^{* * *} \\ & (0.0001) \end{aligned}$ |
| Mother alive (lagged) |  |  | $\begin{aligned} & 0.0118^{* *} \\ & (0.0052) \end{aligned}$ |
| Father alive (lagged) |  |  | $\begin{aligned} & 0.0256^{* * *} \\ & (0.0047) \end{aligned}$ |
| Man | $\begin{aligned} & 12.4184^{* * *} \\ & (0.2215) \end{aligned}$ | $\begin{aligned} & 13.7844 * * * \\ & (0.2459) \end{aligned}$ | $\begin{aligned} & 12.8195^{* * *} \\ & (0.2287) \end{aligned}$ |
| Age | $\begin{aligned} & 7.6209 * * * \\ & (0.1505) \end{aligned}$ | $\begin{aligned} & 8.4592 * * * \\ & (0.1671) \end{aligned}$ | $\begin{aligned} & 7.8671^{* * *} \\ & (0.1554) \end{aligned}$ |
| Age^2 | $\begin{aligned} & -0.0473^{* * *} \\ & (0.0011) \end{aligned}$ | $\begin{aligned} & -0.0525^{* * *} \\ & (0.0012) \end{aligned}$ | $\begin{aligned} & -0.0488^{* * *} \\ & (0.0011) \end{aligned}$ |
| Austria | $\begin{aligned} & 8.3762^{* * *} \\ & (0.4138) \end{aligned}$ | $\begin{aligned} & 9.2976 * * * \\ & (0.4593) \end{aligned}$ | $\begin{aligned} & 8.6468^{* * *} \\ & (0.4272) \end{aligned}$ |
| Germany | $\begin{aligned} & 2.1708^{* * *} \\ & (0.4249) \end{aligned}$ | $\begin{aligned} & 2.4096^{* * *} \\ & (0.4716) \end{aligned}$ | $\begin{aligned} & 2.2409 * * * \\ & (0.4386) \end{aligned}$ |
| Sweden | $\begin{aligned} & 8.3323^{* * *} \\ & (0.4486) \end{aligned}$ | $\begin{aligned} & 9.2489 * * * \\ & (0.4979) \end{aligned}$ | $\begin{aligned} & 8.6014^{* * *} \\ & (0.4631) \end{aligned}$ |
| Spain | $\begin{aligned} & 15.0797 * * * \\ & (0.4282) \end{aligned}$ | $\begin{aligned} & 16.7385 * * * \\ & (0.4753) \end{aligned}$ | $\begin{aligned} & 15.5668^{* * *} \\ & (0.4420) \end{aligned}$ |
| Italy | $\begin{aligned} & 9.3638^{* * *} \\ & (0.4176) \end{aligned}$ | $\begin{aligned} & 10.3938 * * * \\ & (0.4635) \end{aligned}$ | $\begin{aligned} & 9.6663 * * * \\ & (0.4311) \end{aligned}$ |
| France | $\begin{aligned} & -0.4362 \\ & (0.4006) \end{aligned}$ | $\begin{aligned} & -0.4842 \\ & (0.4447) \end{aligned}$ | $\begin{aligned} & -0.4503 \\ & (0.4135) \end{aligned}$ |
| Denmark | $\begin{aligned} & 16.4724^{* * *} \\ & (0.4563) \end{aligned}$ | $\begin{aligned} & 18.2844^{* * *} \\ & (0.5065) \end{aligned}$ | $\begin{aligned} & 17.0045 * * * \\ & (0.4710) \end{aligned}$ |
| Switzerland | $\begin{aligned} & 3.7549^{* * *} \\ & (0.4498) \end{aligned}$ | $\begin{aligned} & 4.1679 * * * \\ & (0.4993) \end{aligned}$ | $\begin{aligned} & 3.8762^{* * *} \\ & (0.4643) \end{aligned}$ |
| Wave 1-2 | $\begin{aligned} & -2.0010^{* * *} \\ & (0.3004) \end{aligned}$ | $\begin{aligned} & -2.2211^{* * *} \\ & (0.3334) \end{aligned}$ | $\begin{aligned} & -2.0656 * * * \\ & (0.3101) \end{aligned}$ |
| Wave 2-4 | $\begin{aligned} & -30.2523 * * * \\ & (0.2982) \end{aligned}$ | $\begin{aligned} & -33.5801^{* * *} \\ & (0.3310) \end{aligned}$ | $\begin{aligned} & -31.2294^{* * *} \\ & (0.3078) \end{aligned}$ |
| Wave 4-5 | $\begin{aligned} & 3.0392 * * * \\ & (0.3303) \end{aligned}$ | $\begin{aligned} & 3.3735 * * * \\ & (0.3666) \end{aligned}$ | $\begin{aligned} & 3.1374 * * * \\ & (0.3410) \end{aligned}$ |
| No education | $\begin{aligned} & 10.0616^{* * *} \\ & (0.5731) \end{aligned}$ | $\begin{aligned} & 11.1684^{* * *} \\ & (0.6361) \end{aligned}$ | $\begin{aligned} & 10.3866 * * * \\ & (0.5916) \end{aligned}$ |
| Primary education | $\begin{aligned} & 11.4072^{* * *} \\ & (0.3523) \end{aligned}$ | $\begin{aligned} & 12.6620^{* * *} \\ & (0.3911) \end{aligned}$ | $\begin{aligned} & 11.7757 * * * \\ & (0.3637) \end{aligned}$ |
| Lower secondary education | $\begin{aligned} & 5.3111^{* * *} \\ & (0.3527) \end{aligned}$ | $\begin{aligned} & 5.8953^{* * *} \\ & (0.3915) \end{aligned}$ | $\begin{aligned} & 5.4826^{* * *} \\ & (0.3641) \end{aligned}$ |
| Upper secondary education | $\begin{aligned} & 8.8285^{* * *} \\ & (0.2856) \end{aligned}$ | $\begin{aligned} & 9.7996^{* * *} \\ & (0.3170) \end{aligned}$ | $\begin{aligned} & 9.1137 * * * \\ & (0.2948) \end{aligned}$ |
| Post-secondary non-tertiary | $\begin{aligned} & 11.4665 * * * \\ & (0.5977) \end{aligned}$ | $\begin{aligned} & 12.7278^{* * *} \\ & (0.6634) \end{aligned}$ | $\begin{aligned} & 11.8369 * * * \\ & (0.6170) \end{aligned}$ |
| Married/cohabiting | $\begin{aligned} & -1.0799 * * \\ & (0.4435) \end{aligned}$ | $\begin{aligned} & -1.1987 * * * \\ & (0.4923) \end{aligned}$ | $\begin{aligned} & -1.1148^{* * *} \\ & (0.4578) \end{aligned}$ |
| Single | $\begin{aligned} & -0.7104 \\ & (0.5236) \end{aligned}$ | $\begin{aligned} & -0.7885 \\ & (0.5812) \end{aligned}$ | $\begin{aligned} & -0.7333 \\ & (0.5405) \end{aligned}$ |
| Separated/Divorced | $\begin{aligned} & 1.7195^{* * *} \\ & (0.4628) \end{aligned}$ | $\begin{aligned} & 1.9086^{* * *} \\ & (0.5137) \end{aligned}$ | $\begin{aligned} & 1.7750^{* * *} \\ & (0.4777) \end{aligned}$ |
| Days in hospital (surgery/medical tests) | $\begin{aligned} & -0.1797 * * * \\ & (0.0138) \end{aligned}$ | $\begin{aligned} & -0.1995 * * * \\ & (0.0153) \end{aligned}$ | $\begin{aligned} & -0.1855 * * * \\ & (0.0142) \end{aligned}$ |
| Days in hospital (mental health) | $\begin{aligned} & -0.1120^{* * *} \\ & (0.0541) \end{aligned}$ | $\begin{aligned} & -0.1243^{* * *} \\ & (0.0601) \end{aligned}$ | $\begin{aligned} & -0.1156 * * * \\ & (0.0558) \end{aligned}$ |
| Days in other institutions | $\begin{aligned} & 0.0642^{* * *} \\ & (0.0209) \end{aligned}$ | $\begin{aligned} & 0.0713^{* * *} \\ & (0.0232) \end{aligned}$ | $\begin{aligned} & 0.0663^{* * *} \\ & (0.0216) \end{aligned}$ |
| Vistis to GP | $\begin{aligned} & -0.0864^{* * *} \\ & (0.0223) \end{aligned}$ | $\begin{aligned} & -0.0959 * * * \\ & (0.0248) \end{aligned}$ | $\begin{aligned} & -0.0892 * * * \\ & (0.0230) \end{aligned}$ |
| Employed | $\begin{aligned} & 0.7910^{* *} \\ & (0.3805) \end{aligned}$ | $\begin{aligned} & 0.8780^{* *} \\ & (0.4224) \end{aligned}$ | $\begin{aligned} & 0.8165^{* *} \\ & (0.3928) \end{aligned}$ |
| Unemployed | $\begin{aligned} & 1.1400^{*} \\ & (0.6498) \end{aligned}$ | $\begin{aligned} & 1.2654^{*} \\ & (0.7213) \end{aligned}$ | $\begin{aligned} & 1.1768^{*} \\ & (0.6708) \end{aligned}$ |
| Retired | $\begin{aligned} & 1.3713^{* * *} \\ & (0.3353) \end{aligned}$ | $\begin{aligned} & 1.5221^{* * *} \\ & (0.3722) \end{aligned}$ | $\begin{aligned} & 1.4156^{* * *} \\ & (0.3461) \end{aligned}$ |
| Lives alone | $\begin{aligned} & -2.4425 * * * \\ & (0.3972) \end{aligned}$ | $\begin{aligned} & -2.7112 * * * \\ & (0.4409) \end{aligned}$ | $\begin{aligned} & -2.5214^{* * *} \\ & (0.4100) \end{aligned}$ |
| Adjusted income (1,000 PPP) | $\begin{aligned} & 66.417^{* *} \\ & (30.565) \end{aligned}$ | $\begin{aligned} & 73.7229 * * \\ & (33.9272) \end{aligned}$ | $\begin{aligned} & 68.5623^{* *} \\ & (31.5522) \end{aligned}$ |
| Adjusted wealth (1,000 PPP) | $\begin{aligned} & 1.1664^{* * *} \\ & (0.3838) \end{aligned}$ | $\begin{aligned} & 1.2947 * * * \\ & (0.4260) \end{aligned}$ | $\begin{aligned} & 1.2041^{* * *} \\ & (0.3962) \end{aligned}$ |


| Rural área, small village | $-0.8729^{* * *}$ | $-0.9689^{* * *}$ | $-0.9011^{* * *}$ |
| :--- | :--- | :--- | :--- |
|  | $(0.2135)$ | $(0.2370)$ | $(0.2204)$ |
| Constant | $-29.6701^{* * *}$ | $-32.9338^{* * *}$ | $-30.6284^{* * *}$ |
|  | $(5.532)$ | $(6.1405)$ | 17.2107 |
| N | 69,990 | 0.175 | 17,223 |
| $\mathrm{R}^{2}$ | 0.3825 | 129.828 | 0.176 |
| F | 916.898 | 0.000 | 130.924 |
| p | 0.0000 | 0.000 |  |

Omitted variables: women, tertiary education, widow, relation with economic activity: houseworking and other situations, Belgium, waves 5-6. Bootstrap with 1,000 replications. Robust standard errors between parenthesis. * statistically significant at $10 \%$; ** statistically significant at $5 \%$; *** statistically significant at $1 \%$. Income and wealth are expressed in 1,000 units of Purchase Power of Parity and adjusted by household sized (dividing wealth and income by the square root of the number of household members).
${ }^{\text {a }}$ Recorded father's age at decease if father has died at the time of the survey. Predicted father's age at decease if father is alive at the time of the survey.
${ }^{\text {a }}$ Recorded mother's age at decease if mother has died at the time of the survey. Predicted mother's age at decease if mother is alive at the time of the survey.

## Appendix E

## Plausible exogeneity of the instruments:

The departure point is equation (2) in which we explicitly distinguish between the potential endogenous variable $\left(B S E_{i c t}\right)$ and the other explanatory variables:

$$
\begin{equation*}
I N D_{i c t}=\alpha_{0}+\alpha_{1} B S E_{i c t}+\alpha_{2} X_{i c t}+T_{t}+C_{c}+\epsilon_{i c t} \tag{C.1}
\end{equation*}
$$

The first method is the $\gamma$-Local-to-Zero (LTZ) approximation bounds method, which introduces some bias term (or exogeneity error) in the approximate distribution of $\widehat{\alpha_{1}}$. In other words, it relaxes the exclusion restriction requirement by allowing for uncertainty in the priors about $\gamma$. According to Conley et al., (2012) this method provides robustness with respect to 2SLS approach under the assumption that the priors are correct.

$$
\begin{align*}
& \widehat{\alpha_{1}} \sim N\left(\lambda, \Sigma_{2 S L S}\right)+\Pi \gamma  \tag{C.2}\\
& \gamma \sim \Upsilon \\
& \Pi=\left(X^{\prime} Z\left(Z^{\prime} Z\right)^{-1} Z^{\prime} X\right)^{-1} X^{\prime} Z
\end{align*}
$$

where $\Upsilon$ is the distribution of $\gamma, \Sigma_{2 S L S}$ is the variance-covariance matrix for the estimation 2SLS and $Z$ is the vector of instrumental variables. The distribution of the exogeneity error ( $\widehat{\Pi \gamma}$ ) depends on the sample moments of the matrix $\Pi$, which shows a negative relationship between the strength of the instrumental variable and the exogeneity error, and the distribution $\Upsilon$. This exogeneity error is an indicator of the deviations of $\widehat{\alpha_{1}}$ from the asymptotic standard distribution of the 2SLS estimator due to non-fulfilment of the exclusion restriction assumption.
It is assumed that $\gamma$ follows a normal distribution with mean $\mu_{\gamma}$ and variance-covariance matrix $\Omega_{\gamma}$. Then, the asymptotic distribution $\widehat{\alpha_{1}}$ of can be expressed as:

$$
\begin{equation*}
\widehat{\alpha_{1}} \sim N\left(\lambda+\Pi \mu_{\gamma}, \Sigma_{2 S L S}+\Pi \Omega_{\gamma} \Pi^{\prime}\right) \tag{С.3}
\end{equation*}
$$

Following Conley et al. (2012), we implement the simples form of priors for $\gamma$, that is, $\gamma \sim N\left(0, \delta^{2}\right)$ and computed the $95 \%$ confidence intervals for $\alpha_{1}$ for different values of $\delta$. Under the assumption that priors are correct, this approach provides valid inference and robustness with respect to normal 2SLS approach.

The second method is the Union Confidence Interval (UCI), which allows us to analyse the robustness of the estimations in case of a direct relationship between the instrumental variables (for notation simplicity, parent's age of decease and parent's living status have been collapsed in the variable, $P A D \& L S_{i t}$ ) and the outcome variables. Following Conley et al. (2012) equation (C.1) can be modified as follows:

$$
\begin{align*}
& I N D_{i c t}=\alpha_{0}+\alpha_{1} B S E_{i c t}+\alpha_{2} X_{i c t}+\alpha_{3} \text { PAD } L S_{i c t}+T_{t}+C_{c}+\epsilon_{i c t}  \tag{C.4}\\
& B S E_{i c t}=\beta_{0}+\alpha_{3} P_{A D} L L S_{i c t}++\beta_{5} X_{i c t}+T_{t}+C_{c}+\varepsilon_{i c t}
\end{align*}
$$

In a normal 2SLS estimation the term ( $\alpha_{3} P A D_{i c t}$ ) would not be present in equation (C.4). If the strict exogeneity assumption is satisfied, parents' age of decease does not have any effect over outcome variables and thus $\alpha_{3}=0$. The innovation proposed by Conley et a. (2012) consist in relaxing the strict exogeneity assumption ( $\alpha_{3} \neq 0$ ) and checking its significance in the outcome equation. Then, allowing for non-zero $\gamma$, equation (C.4) can be expressed as (C.5)

Considering that the outcome variable is now ( $I N D_{i c t}-\alpha_{3} P A D \& L S_{i c t}$ ), then $\alpha_{1}$ can be consistently estimated using PAD as an instrument for $B S E$. Under the UCI approach, $\alpha_{1}$ is estimated given any $\alpha_{3}$ belonging to the specific support interval for $\alpha_{3}: \alpha_{3} \in[-\delta,+\delta]$. Conley et al. (2012) notes that given that $\gamma$ belongs to that interval, the union will contain the true parameter value for $\alpha_{1}$ at least $95 \%$ of the time (if using a $95 \%$ confidence interval).

Figure E1. Local-to-Zero approximation and Union of Confidence Intervals. Test for instrument validity for the instrument "father's age of decease". ${ }^{53}$


[^30]Table E1. Effect of lagged BSE and of the death of a brother, sister or child between the current wave and the previous one

|  | M1 | M2 | M3 | M4 | M5 | M6 | M7 | M8 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| BSE (lagged) | $0.3048^{* * *}$ (0.0073) | $0.2962^{* * *}$ <br> (0.0073) | $0.2719^{* * *}$ <br> (0.0072) | $0.2714^{* * *}$ (0.0072) | $0.2714^{* * *}$ <br> (0.0072) | $0.2714^{* * *}$ (0.0072) | $0.2713^{* * *}$ <br> (0.0072) | $0.2714^{* * *}$ (0.0072) |
| Brother dies between waves |  |  |  |  | $\begin{aligned} & -0.2758 \\ & (0.6714) \end{aligned}$ |  |  | $\begin{aligned} & 0.0055 \\ & (0.6947) \end{aligned}$ |
| Sister dies between waves |  |  |  |  |  | $\begin{aligned} & -0.1086^{*} \\ & (0.0615) \end{aligned}$ |  | $\begin{gathered} -0.0842 \\ (0.0666) \end{gathered}$ |
| Child dies between waves |  |  |  |  |  |  | $\begin{aligned} & 0.3828 \\ & (1.555) \end{aligned}$ | $\begin{aligned} & 0.4084 \\ & (1.556) \end{aligned}$ |
| Man | $\begin{aligned} & 8.4185^{* * *} \\ & (0.4157) \end{aligned}$ | $\begin{aligned} & 8.7439 * * * \\ & (0.4266) \end{aligned}$ | $\begin{aligned} & 8.1307^{* * *} \\ & (0.4273) \end{aligned}$ | $\begin{aligned} & 7.8798^{* * *} \\ & (0.4429) \end{aligned}$ | $\begin{aligned} & 7.8813^{* * *} \\ & \text { (0.4429) } \end{aligned}$ | $\begin{aligned} & 7.8722^{* * *} \\ & \text { (0.4429) } \end{aligned}$ | $\begin{aligned} & 7.8773^{* * *} \\ & (0.4429) \end{aligned}$ | $\begin{aligned} & 7.8695^{* * *} \\ & (0.4430) \end{aligned}$ |
| Age | $\begin{aligned} & 2.3918^{* * *} \\ & (0.3104) \end{aligned}$ | $2.3516^{* * *}$ <br> (0.3106) | $2.4590^{* * *}$ <br> (0.3070) | $\begin{aligned} & 2.1123^{* * *} \\ & (0.3472) \end{aligned}$ | $\begin{aligned} & 2.1132^{* * *} \\ & (0.3472) \end{aligned}$ | $2.1090^{* * *}$ <br> (0.3472) | $\begin{aligned} & 2.1145^{* * *} \\ & (0.3473) \end{aligned}$ | $\begin{aligned} & 2.1114^{* * *} \\ & (0.3473) \end{aligned}$ |
| Age^2 | $\begin{aligned} & -0.0136^{* * *} \\ & (0.0022) \end{aligned}$ | $\begin{aligned} & -0.0135^{* * *} \\ & (0.0022) \end{aligned}$ | $\begin{aligned} & -0.0138^{* * *} \\ & (0.0022) \end{aligned}$ | $\begin{aligned} & -0.0116^{* * *} \\ & (0.0024) \end{aligned}$ | $\begin{aligned} & -0.0116^{* * *} \\ & (0.0024) \end{aligned}$ | $\begin{aligned} & -0.0115^{* * *} \\ & (0.0024) \end{aligned}$ | $\begin{aligned} & -0.0116^{* * *} \\ & (0.0024) \end{aligned}$ | $\begin{aligned} & -0.0115^{* * *} \\ & (0.0024) \end{aligned}$ |
| Austria | $\begin{aligned} & 6.1510^{* * *} \\ & (0.9568) \end{aligned}$ | $\begin{aligned} & 5.4431^{* * *} \\ & (0.9603) \end{aligned}$ | $\begin{aligned} & 5.6171^{* * *} \\ & (0.9441) \end{aligned}$ | $\begin{aligned} & 5.4256 * * * \\ & (0.9459) \end{aligned}$ | $\begin{aligned} & 5.4319^{* * *} \\ & (0.9461) \end{aligned}$ | $\begin{aligned} & 5.4321^{* * *} \\ & (0.9459) \end{aligned}$ | $\begin{aligned} & 5.4187^{* * *} \\ & (0.9462) \end{aligned}$ | $\begin{aligned} & 5.4247^{* * *} \\ & (0.9463) \end{aligned}$ |
| Germany | 0.2639 $(0.7875)$ | 0.5996 (0.7993) | 1.4964* (0.7898) | 1.4155* (0.7928) | 1.4199* (0.7929) | 1.4396* (0.7929) | 1.4133* (0.7928) | 1.4373* (0.7930) |
| Sweden | 5.9098*** | 5.5205*** | 4.3033*** | 4.0283*** | 4.0290*** | 4.0295*** | 4.0212*** | 4.0220*** |
|  | (0.7524) | (0.7578) | (0.7504) | (0.7695) | (0.7696) | (0.7695) | (0.7698) | (0.7698) |
| Spain | $\begin{aligned} & 9.9827^{* * *} \\ & (0.8187) \end{aligned}$ | $\begin{aligned} & 8.5335 * * * \\ & (0.8651) \end{aligned}$ | $\begin{aligned} & 9.7920^{* * *} \\ & (0.8539) \end{aligned}$ | $\begin{aligned} & 9.6527^{* * *} \\ & (0.8643) \end{aligned}$ | $\begin{aligned} & 9.6686^{* * *} \\ & \text { (0.8652) } \end{aligned}$ | $\begin{aligned} & 9.7147^{* * *} \\ & \text { (0.8651) } \end{aligned}$ | $\begin{aligned} & 9.6555^{* * *} \\ & \text { (0.8644) } \end{aligned}$ | $\begin{aligned} & 9.7177^{* * *} \\ & (0.8657) \end{aligned}$ |
| Italy | $\begin{aligned} & 8.4141^{* * *} \\ & (0.7087) \end{aligned}$ | $\begin{aligned} & 6.5027^{* * *} \\ & (0.7320) \end{aligned}$ | $\begin{aligned} & 7.9287^{* * *} \\ & (0.7253) \end{aligned}$ | $\begin{aligned} & \text { 7.7671*** } \\ & (0.7268) \end{aligned}$ | $\begin{aligned} & 7.7712^{* * *} \\ & (0.7269) \end{aligned}$ | $\begin{aligned} & 7.7835^{* * *} \\ & (0.7268) \end{aligned}$ | $\begin{aligned} & 7.7722^{* * *} \\ & (0.7269) \end{aligned}$ | $\begin{aligned} & \text { 7.7889*** } \\ & (0.7270) \end{aligned}$ |
| France | $\begin{aligned} & 1.2744^{*} \\ & (0.7710) \end{aligned}$ | $\begin{aligned} & 0.8496 \\ & (0.7810) \end{aligned}$ | $\begin{aligned} & 11.326 \\ & (0.7687) \end{aligned}$ | $\begin{aligned} & 0.8813 \\ & (0.7739) \end{aligned}$ | $\begin{aligned} & 0.8820 \\ & (0.7739) \end{aligned}$ | $\begin{aligned} & 0.8932 \\ & (0.7739) \end{aligned}$ | $\begin{aligned} & 0.8798 \\ & (0.7739) \end{aligned}$ | $\begin{aligned} & 0.8917 \\ & (0.7739) \end{aligned}$ |
| Denmark | $\begin{aligned} & 11.2501^{* * *} \\ & (0.7640) \end{aligned}$ | $\begin{aligned} & 11.5870^{* * *} \\ & (0.7703) \end{aligned}$ | $\begin{aligned} & 10.4852^{* * *} \\ & (0.7658) \end{aligned}$ | $\begin{aligned} & 10.3882^{* * *} \\ & (0.7707) \end{aligned}$ | $\begin{aligned} & 10.3900^{* * *} \\ & (0.7707) \end{aligned}$ | $\begin{aligned} & 10.4066^{* * *} \\ & (0.7707) \end{aligned}$ | $\begin{aligned} & 10.3844^{* * *} \\ & (0.7708) \end{aligned}$ | $\begin{aligned} & 10.4026^{* * *} \\ & (0.7708) \end{aligned}$ |
| Switzerland | $\begin{aligned} & 5.0121^{* * *} \\ & (0.8835) \end{aligned}$ | $\begin{aligned} & 3.8189 * * * \\ & (0.9137) \end{aligned}$ | $\begin{aligned} & 2.8765^{* * *} \\ & (0.9016) \end{aligned}$ | $\begin{aligned} & 3.0333^{* * *} \\ & (0.9088) \end{aligned}$ | $\begin{aligned} & 3.0397^{* * *} \\ & (0.9089) \end{aligned}$ | $\begin{aligned} & 3.0357^{* * *} \\ & (0.9087) \end{aligned}$ | $\begin{aligned} & 3.0309 * * * \\ & (0.9088) \end{aligned}$ | $\begin{aligned} & 3.0331^{* * *} \\ & (0.9089) \end{aligned}$ |
| Wave 2-4 | $\begin{aligned} & -42.2337^{* * *} \\ & (0.4844) \end{aligned}$ | $\begin{aligned} & -42.3502^{* * *} \\ & (0.4848) \end{aligned}$ | $\begin{aligned} & -41.4680^{* * *} \\ & (0.4782) \end{aligned}$ | $\begin{aligned} & -41.5568^{* *} \\ & (0.4794) \end{aligned}$ | $\begin{aligned} & -41.5692^{* *} \\ & (0.4803) \end{aligned}$ | $\begin{aligned} & -41.5806^{* *} \\ & (0.4796) \end{aligned}$ | $\begin{aligned} & -41.5539^{* * *} \\ & (0.4794) \end{aligned}$ | $\begin{aligned} & -41.5773^{* * *} \\ & (0.4804) \end{aligned}$ |
| Wave 4-5 | $\begin{gathered} -0.9282 \\ (33.982) \end{gathered}$ | $\begin{aligned} & -0.4881 \\ & (33.825) \end{aligned}$ | $\begin{aligned} & -0.7728 \\ & (33.224) \end{aligned}$ | $\begin{aligned} & -0.5758 \\ & (33.254) \end{aligned}$ | $\begin{aligned} & -0.5134 \\ & (33.290) \end{aligned}$ | $\begin{gathered} -0.3422 \\ (33.283) \end{gathered}$ | $\begin{gathered} -0.6065 \\ (33.266) \end{gathered}$ | $\begin{gathered} -0.3751 \\ (33.316) \end{gathered}$ |
| No education |  | $\begin{aligned} & 4.1719^{* * *} \\ & (1.913) \end{aligned}$ | $\begin{aligned} & 6.6472^{* * *} \\ & (1.834) \end{aligned}$ | $\begin{aligned} & 6.9881^{* * *} \\ & (1.920) \end{aligned}$ | $\begin{aligned} & 7.0027^{* * *} \\ & (1.926) \end{aligned}$ | $\begin{aligned} & 7.0241^{* * *} \\ & (1.921) \end{aligned}$ | $\begin{aligned} & 6.9881^{* * *} \\ & (1.920) \end{aligned}$ | $\begin{aligned} & 7.0241^{* * *} \\ & (1.926) \end{aligned}$ |
| Primary education |  | $\begin{aligned} & 7.6616^{* * *} \\ & (0.6513) \end{aligned}$ | $\begin{aligned} & 9.4152^{* * *} \\ & (0.6514) \end{aligned}$ | $\begin{aligned} & 9.6512^{* * *} \\ & (0.6610) \end{aligned}$ | $\begin{aligned} & 9.6626^{* * *} \\ & (0.6616) \end{aligned}$ | $\begin{aligned} & 9.6973^{* * *} \\ & (0.6615) \end{aligned}$ | $\begin{aligned} & 9.6509^{* * *} \\ & (0.6610) \end{aligned}$ | $\begin{aligned} & 9.6969 * * * \\ & (0.6619) \end{aligned}$ |
| Lower secondary education |  | $\begin{aligned} & 2.6195^{* * *} \\ & (0.6775) \end{aligned}$ | $\begin{aligned} & 3.6567^{* * *} \\ & (0.6701) \end{aligned}$ | $\begin{aligned} & 3.7777^{* * *} \\ & (0.6747) \end{aligned}$ | $\begin{aligned} & 3.7816^{* * *} \\ & (0.6748) \end{aligned}$ | $\begin{aligned} & 3.7905^{* * *} \\ & (0.6747) \end{aligned}$ | $\begin{aligned} & 3.7764^{* * *} \\ & (0.6747) \end{aligned}$ | $\begin{aligned} & 3.7891^{* * *} \\ & (0.6748) \end{aligned}$ |
| Upper secondary education |  | $\begin{aligned} & 5.2170^{* * *} \\ & (0.5697) \end{aligned}$ | $\begin{aligned} & 6.2601^{* * *} \\ & (0.5638) \end{aligned}$ | $\begin{aligned} & 6.3199 * * * \\ & (0.5661) \end{aligned}$ | $\begin{aligned} & 6.3218^{* * *} \\ & (0.5662) \end{aligned}$ | $\begin{aligned} & 6.3302^{* * *} \\ & (0.5661) \end{aligned}$ | $\begin{aligned} & 6.3202^{* * *} \\ & (0.5661) \end{aligned}$ | $\begin{aligned} & 6.3305^{* * *} \\ & (0.5662) \end{aligned}$ |
| Post-secondary non-tertiary |  | $\begin{aligned} & 7.1224^{* * *} \\ & (1.176) \end{aligned}$ | $\begin{aligned} & 7.3119^{* * *} \\ & (1.964) \end{aligned}$ | $\begin{aligned} & \text { 7.4809*** } \\ & \text { (1.971) } \end{aligned}$ | $\begin{aligned} & 7.4795^{* * *} \\ & (1.971) \end{aligned}$ | $\begin{aligned} & 7.4893^{* * *} \\ & (1.970) \end{aligned}$ | $\begin{aligned} & 7.4763^{* * *} \\ & (1.972) \end{aligned}$ | $\begin{aligned} & 7.4845^{* * *} \\ & (1.972) \end{aligned}$ |
| Married/cohabiting |  | $\begin{aligned} & 0.3161 \\ & (0.6284) \end{aligned}$ | $\begin{aligned} & -0.0428 \\ & (0.6194) \end{aligned}$ | $\begin{aligned} & -1.5270^{*} \\ & (0.8232) \end{aligned}$ | $\begin{aligned} & -1.5268^{*} \\ & (0.8232) \end{aligned}$ | $\begin{aligned} & -1.5156^{*} \\ & (0.8232) \end{aligned}$ | $\begin{aligned} & -1.5197^{*} \\ & (0.8235) \end{aligned}$ | $\begin{aligned} & -1.5078^{*} \\ & (0.8235) \end{aligned}$ |
| Single |  | $\begin{aligned} & 0.5703 \\ & (10.376) \end{aligned}$ | $\begin{aligned} & 0.4080 \\ & (10.205) \end{aligned}$ | $\begin{aligned} & 0.1976 \\ & (10.222) \end{aligned}$ | $\begin{aligned} & 0.2069 \\ & (10.225) \end{aligned}$ | $\begin{aligned} & 0.2223 \\ & (10.223) \end{aligned}$ | $\begin{aligned} & 0.2101 \\ & (10.228) \end{aligned}$ | $\begin{aligned} & 0.2355 \\ & (10.231) \end{aligned}$ |
| Separated/Divorced |  | $\begin{aligned} & 0.9825 \\ & (0.9035) \end{aligned}$ | $\begin{aligned} & 0.8896 \\ & (0.8879) \end{aligned}$ | $\begin{aligned} & 0.6562 \\ & (0.8908) \end{aligned}$ | $\begin{aligned} & 0.6582 \\ & (0.8909) \end{aligned}$ | $\begin{aligned} & 0.6682 \\ & (0.8908) \end{aligned}$ | $\begin{aligned} & 0.6592 \\ & (0.8909) \end{aligned}$ | $\begin{aligned} & 0.6713 \\ & (0.8909) \end{aligned}$ |
| Depressed |  |  | $\begin{aligned} & -1.8669^{* * *} \\ & (0.4308) \end{aligned}$ | $\begin{aligned} & -1.8393^{* * *} \\ & (0.4310) \end{aligned}$ | $\begin{aligned} & -1.8384^{* * *} \\ & (0.4310) \end{aligned}$ | $\begin{aligned} & -1.8281^{* * *} \\ & (0.4310) \end{aligned}$ | $\begin{aligned} & -1.8444^{* * *} \\ & (0.4312) \end{aligned}$ | $\begin{aligned} & -1.8334^{* * *} \\ & (0.4313) \end{aligned}$ |
| Self-reported health: excellent |  |  | $\begin{aligned} & 13.6700^{* * *} \\ & (1.726) \end{aligned}$ | $\begin{aligned} & 13.5664^{* * *} \\ & (1.771) \end{aligned}$ | $\begin{aligned} & 13.5670^{* * *} \\ & (1.771) \end{aligned}$ | $\begin{aligned} & 13.5518^{* * *} \\ & (1.771) \end{aligned}$ | $\begin{aligned} & 13.5679 * * * \\ & (1.772) \end{aligned}$ | $\begin{aligned} & 13.5533^{* * *} \\ & (1.772) \end{aligned}$ |
| Self-reported health: very good |  |  | $\begin{aligned} & 9.8517^{* * *} \\ & (0.9481) \end{aligned}$ | $\begin{aligned} & 9.7131^{* * *} \\ & (0.9521) \end{aligned}$ | $\begin{aligned} & 9.7122^{* * *} \\ & (0.9522) \end{aligned}$ | $\begin{aligned} & 9.6850 * * \\ & (0.9522) \end{aligned}$ | $\begin{aligned} & 9.7178^{* * *} \\ & (0.9523) \end{aligned}$ | $\begin{aligned} & 9.6899 * * * \\ & (0.9524) \end{aligned}$ |
| Self-reported health: good |  |  | $\begin{aligned} & 6.4866^{* * *} \\ & (0.8493) \end{aligned}$ | $\begin{aligned} & 6.3695^{* * *} \\ & (0.8524) \end{aligned}$ | $\begin{aligned} & 6.3686 * * * \\ & (0.8524) \end{aligned}$ | $\begin{aligned} & 6.3553^{* * *} \\ & (0.8524) \end{aligned}$ | $\begin{aligned} & 6.3740^{* * *} \\ & (0.8525) \end{aligned}$ | $\begin{aligned} & 6.3600^{* * *} \\ & (0.8525) \end{aligned}$ |
| Self-reported health: fair |  |  | $\begin{aligned} & 3.1174^{* * *} \\ & (0.8369) \end{aligned}$ | $\begin{aligned} & 3.0339 * * * \\ & (0.8379) \end{aligned}$ | $\begin{aligned} & 3.0343^{* * *} \\ & (0.8379) \end{aligned}$ | $\begin{aligned} & 3.0203^{* * *} \\ & (0.8379) \end{aligned}$ | $\begin{aligned} & 3.0381^{* * *} \\ & (0.8380) \end{aligned}$ | $\begin{aligned} & 3.0247^{* * *} \\ & (0.8380) \end{aligned}$ |
| Hypertension |  |  | $\begin{aligned} & -3.2354^{* * *} \\ & (0.4243) \end{aligned}$ | $\begin{aligned} & -3.2456^{* * *} \\ & (0.4241) \end{aligned}$ | $\begin{aligned} & -3.2460^{* * *} \\ & (0.4241) \end{aligned}$ | $\begin{aligned} & -3.2568^{* * *} \\ & (0.4242) \end{aligned}$ | $\begin{aligned} & -3.2441^{* * *} \\ & (0.4242) \end{aligned}$ | $\begin{aligned} & -3.2553^{* * *} \\ & (0.4242) \end{aligned}$ |
| Cancer |  |  | $\begin{aligned} & 15.7267^{* * *} \\ & (1.575) \end{aligned}$ | $\begin{aligned} & 15.6632^{* * *} \\ & (1.572) \end{aligned}$ | $\begin{aligned} & 15.6615^{* * *} \\ & (1.573) \end{aligned}$ | $\begin{aligned} & 15.6336^{* * *} \\ & (1.573) \end{aligned}$ | $\begin{aligned} & 15.6688^{* * *} \\ & (1.574) \end{aligned}$ | $\begin{aligned} & 15.6394^{* * *} \\ & (1.575) \end{aligned}$ |
| Stroke |  |  | $\begin{aligned} & 7.3373^{* * *} \\ & (1.756) \end{aligned}$ | $\begin{aligned} & 7.3578^{* * *} \\ & (1.755) \end{aligned}$ | $\begin{aligned} & 7.3594^{* * *} \\ & (1.755) \end{aligned}$ | $\begin{aligned} & 7.3499 * * \\ & (1.754) \end{aligned}$ | $\begin{aligned} & 7.3634^{* * *} \\ & (1.756) \end{aligned}$ | $\begin{aligned} & 7.3558^{* * *} \\ & (1.756) \end{aligned}$ |
| Diabetes |  |  | $\begin{aligned} & 6.2485^{* * *} \\ & (0.6390) \end{aligned}$ | $\begin{aligned} & 6.2237^{* * *} \\ & (0.6389) \end{aligned}$ | $\begin{aligned} & 6.2264^{* * *} \\ & (0.6390) \end{aligned}$ | $\begin{aligned} & 6.2188^{* * *} \\ & (0.6389) \end{aligned}$ | $\begin{aligned} & 6.2181^{* * *} \\ & (0.6391) \end{aligned}$ | $\begin{aligned} & 6.2128^{* * *} \\ & (0.6392) \end{aligned}$ |
| Heart attack |  |  | $\begin{aligned} & -1.7856^{* * *} \\ & (0.6515) \end{aligned}$ | $\begin{aligned} & -1.8176^{* * *} \\ & (0.6513) \end{aligned}$ | $\begin{aligned} & -1.8155^{* * *} \\ & (0.6514) \end{aligned}$ | $\begin{aligned} & -1.8086^{* * *} \\ & (0.6513) \end{aligned}$ | $\begin{aligned} & -1.8175^{* * *} \\ & (0.6514) \end{aligned}$ | $\begin{aligned} & -1.8085^{* * *} \\ & (0.6514) \end{aligned}$ |
| Lung disease |  |  | $\begin{aligned} & 3.9353^{* * *} \\ & (0.8428) \end{aligned}$ | $\begin{aligned} & 3.9327^{* * *} \\ & (0.8425) \end{aligned}$ | $\begin{aligned} & 3.9377^{* * *} \\ & (0.8427) \end{aligned}$ | $\begin{aligned} & 3.9428^{* * *} \\ & (0.8425) \end{aligned}$ | $\begin{aligned} & 3.9310 * * * \\ & (0.8426) \end{aligned}$ | $\begin{aligned} & 3.9410^{* * *} \\ & (0.8427) \end{aligned}$ |
| Alzheimer |  |  | $\begin{aligned} & 31.840 \\ & (21.000) \end{aligned}$ | $\begin{aligned} & 29.371 \\ & (21.008) \end{aligned}$ | $\begin{aligned} & 29.475 \\ & (21.010) \end{aligned}$ | $\begin{aligned} & 29.806 \\ & (21.009) \end{aligned}$ | $\begin{aligned} & 29.369 \\ & (21.009) \end{aligned}$ | $\begin{aligned} & 29.804 \\ & (21.011) \end{aligned}$ |
| Employed |  |  |  | $\begin{aligned} & 0.1741 \\ & (0.7814) \end{aligned}$ | $\begin{aligned} & 0.1724 \\ & (0.7815) \end{aligned}$ | $\begin{aligned} & 0.1687 \\ & (0.7814) \end{aligned}$ | $\begin{aligned} & 0.1706 \\ & (0.7815) \end{aligned}$ | $\begin{aligned} & 0.1650 \\ & (0.7815) \end{aligned}$ |
| Unemployed |  |  |  | $\begin{aligned} & -0.5352 \\ & -14.468 \end{aligned}$ | $\begin{aligned} & -0.5327 \\ & -14.469 \end{aligned}$ | $\begin{aligned} & -0.5288 \\ & -14.468 \end{aligned}$ | $\begin{aligned} & -0.5360 \\ & -14.469 \end{aligned}$ | $\begin{aligned} & -0.5298 \\ & -14.469 \end{aligned}$ |
| Retired |  |  |  | $\begin{aligned} & 1.5448^{* *} \\ & (0.6274) \end{aligned}$ | $\begin{aligned} & 1.5473^{* *} \\ & (0.6275) \end{aligned}$ | $\begin{aligned} & 1.5352^{* *} \\ & (0.6274) \end{aligned}$ | $\begin{aligned} & 1.5459^{* *} \\ & (0.6275) \end{aligned}$ | $\begin{aligned} & 1.5363^{* *} \\ & (0.6275) \end{aligned}$ |
| Lives alone |  |  |  | $\begin{aligned} & -2.0935^{* * *} \\ & (0.7622) \end{aligned}$ | $\begin{aligned} & -2.0951^{* * *} \\ & (0.7623) \end{aligned}$ | $\begin{aligned} & -2.0902^{* * *} \\ & (0.7622) \end{aligned}$ | $\begin{aligned} & -2.0887^{* * *} \\ & (0.7624) \end{aligned}$ | $\begin{aligned} & -2.0851^{* * *} \\ & (0.7624) \end{aligned}$ |
| Rural área, small village |  |  |  | $\begin{aligned} & -0.8875^{* *} \\ & (0.4149) \end{aligned}$ | $\begin{aligned} & -0.8865^{* *} \\ & (0.4150) \end{aligned}$ | $\begin{aligned} & -0.8815^{* *} \\ & (0.4149) \end{aligned}$ | $\begin{aligned} & -0.8869^{* *} \\ & (0.4150) \end{aligned}$ | $\begin{aligned} & -0.8808^{* *} \\ & (0.4150) \end{aligned}$ |


| Constant | $-76.6230^{* * *}$ | $-78.0149^{* * *}$ | $-91.1640^{* * *}$ | $-76.6743^{* * *}$ | $-76.7010^{* * *}$ | $-76.6039^{* * *}$ | $-76.7646^{* * *}$ | $-76.6994^{* * *}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | $(10.130)$ | $(10.910)$ | $(10.990)$ | $(12.464)$ | $(12.468)$ | $(12.458)$ | $(12.492)$ |  |
| N | 69,990 | 69,990 | 69,990 | 69,990 | 69,990 | 69,990 | 69,990 |  |
| $\mathrm{R}^{2}$ | 0.4519 | 0.4575 | 0.4773 | 0.4779 | 0.4779 | 0.4780 |  |  |
| F | 979.283 | 492.569 | 545.935 | 971.620 | 872.181 | 873.400 |  |  |
| p | 0.0000 | 0.0000 | 0.0000 | 0.0000 | 0.4779 |  |  |  |

Omitted variables: women, tertiary education, widow, relation with economic activity: houseworking and other situations, Belgium, waves 5-6. Bootstrap with 1,000 replications. Robust standard errors between parenthesis. * statistically significant at $10 \%$; ** statistically significant at $5 \%$; *** statistically significant at $1 \%$. All regressions include income and wealth are expressed in 1,000 units of Purchase Power of Parity and adjusted by household sized (dividing wealth and income by the square root of the number of household members).

Table E2. Weibull estimations for the discrete-time hazard model including parental characteristics as explanatory variables

|  | Total sample |  | Excluding focal points |  | Excluding erroneous responses to probabilistic question |  | Including only individuals who have answered SHARELIFE |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coef. | Hazard ratio | Coef. | Hazard ratio | Coef. | Hazard ratio | Coef. | Hazard ratio |
| Father alive | $\begin{aligned} & \hline-0.359 \\ & (0.331) \end{aligned}$ | 0.693 | $\begin{aligned} & -0.357 \\ & (0.319) \end{aligned}$ | 0.688 | $\begin{aligned} & \hline-0.356 \\ & (0.308) \end{aligned}$ | 0.683 | $\begin{aligned} & \hline-0.355 \\ & (0.297) \end{aligned}$ | 0.679 |
| Mother alive | $\begin{aligned} & 0.020 \\ & (0.194) \end{aligned}$ | 1.009 | $\begin{aligned} & 0.020 \\ & (0.190) \end{aligned}$ | 0.999 | $\begin{aligned} & 0.020 \\ & (0.186) \end{aligned}$ | 0.989 | $\begin{aligned} & 0.019 \\ & (0.182) \end{aligned}$ | 0.979 |
| Father's age at decease | $\begin{aligned} & -0.060^{* *} \\ & (0.027) \end{aligned}$ | 0.970** | $\begin{aligned} & -0.060^{* *} \\ & (0.028) \end{aligned}$ | 0.972** | $\begin{aligned} & -0.061^{* *} \\ & (0.028) \end{aligned}$ | 0.970** | $\begin{aligned} & -0.061^{* *} \\ & (0.0281) \end{aligned}$ | 0.973** |
| Mother's age at decease | $\begin{aligned} & -0.037^{* *} \\ & (0.017) \end{aligned}$ | 0.990** | $\begin{aligned} & -0.036 * * \\ & (0.017) \end{aligned}$ | 0.991** | $\begin{aligned} & -0.036^{* *} \\ & (0.017) \end{aligned}$ | 0.992** | $\begin{aligned} & -0.037 * * \\ & (0.017) \end{aligned}$ | 0.992** |
| Man | $\begin{aligned} & 0.629 * * * \\ & (0.082) \end{aligned}$ | 1.849*** | $\begin{aligned} & 0.586 * * * \\ & (0.131) \end{aligned}$ | 1.770*** | $\begin{aligned} & 0.602^{* * *} \\ & (0.089) \end{aligned}$ | 1.799*** | $\begin{aligned} & 0.638^{* * *} \\ & (0.089) \end{aligned}$ | 1.859*** |
| Age | $\begin{aligned} & -0.049 \\ & (0.045) \end{aligned}$ | 0.943 | $\begin{aligned} & -0.012 \\ & (0.074) \end{aligned}$ | 0.978 | $\begin{aligned} & -0.045 \\ & (0.051) \end{aligned}$ | 0.947 | $\begin{aligned} & -0.057 \\ & (0.050) \end{aligned}$ | 0.947 |
| Age^2 | $\begin{aligned} & 0.001^{* * *} \\ & (0.000) \end{aligned}$ | 0.991*** | $\begin{aligned} & 0.001 \\ & (0.001) \end{aligned}$ | 0.991 | $\begin{aligned} & 0.001^{* *} \\ & (0.000) \end{aligned}$ | 0.991** | $\begin{aligned} & 0.003^{* * *} \\ & (0.001) \end{aligned}$ | 0.992*** |
| Employed | $\begin{aligned} & -0.649^{* * *} \\ & (0.179) \end{aligned}$ | 0.517*** | $\begin{aligned} & -0.447^{*} \\ & (0.232) \end{aligned}$ | 0.634* | $\begin{aligned} & -0.575^{* * *} \\ & (0.185) \end{aligned}$ | 0.558*** | $\begin{aligned} & -0.682^{* * *} \\ & (0.187) \end{aligned}$ | 0.526*** |
| Lives alone | $\begin{aligned} & 0.192^{* *} \\ & (0.086) \end{aligned}$ | 1.19** | $\begin{aligned} & 0.134 \\ & (0.147) \end{aligned}$ | 1.130 | $\begin{aligned} & 0.233^{* *} \\ & (0.097) \end{aligned}$ | 1.247** | $\begin{aligned} & 0.202 * * \\ & (0.093) \end{aligned}$ | 1.202** |
| Rural area, village, small town | $\begin{aligned} & -0.028 \\ & (0.070) \end{aligned}$ | 0.963 | $\begin{aligned} & -0.175 \\ & (0.115) \end{aligned}$ | 0.833 | $\begin{aligned} & -0.046 \\ & (0.077) \end{aligned}$ | 0.946 | $\begin{aligned} & -0.034 \\ & (0.075) \end{aligned}$ | 0.969 |
| Normal weight | $\begin{aligned} & -0.245^{* * *} \\ & (0.081) \end{aligned}$ | 0.776*** | $\begin{aligned} & -0.261^{* *} \\ & (0.309) \end{aligned}$ | 0.707** | $\begin{aligned} & -0.219^{* *} \\ & (0.089) \end{aligned}$ | 0.798** | $\begin{aligned} & -0.294^{* * *} \\ & (0.084) \end{aligned}$ | 0.795*** |
| High blood pressure | $\begin{aligned} & -0.087^{* *} \\ & (0.036) \end{aligned}$ | 0.908** | $\begin{aligned} & -0.224^{* *} \\ & (0.106) \end{aligned}$ | 0.793** | $\begin{aligned} & -0.109 * * \\ & (0.043) \end{aligned}$ | 0.888** | $\begin{aligned} & -0.091^{* *} \\ & (0.037) \end{aligned}$ | 0.910** |
| Heart attack | $\begin{aligned} & 0.252^{*} \\ & (0.125) \end{aligned}$ | 1.270 | $\begin{aligned} & 0.249^{*} \\ & (0.126) \end{aligned}$ | 1.268 | $\begin{aligned} & 0.240^{*} \\ & (0.117) \end{aligned}$ | 1.249* | $\begin{aligned} & 0.259^{*} \\ & (0.127) \end{aligned}$ | 1.274* |
| Cancer | $\begin{aligned} & 0.934^{* * *} \\ & (0.101) \end{aligned}$ | 2.500*** | $\begin{aligned} & 1.007^{* * *} \\ & (0.155) \end{aligned}$ | 2.689*** | $\begin{aligned} & 0.902^{* * *} \\ & (0.116) \end{aligned}$ | 2.422*** | $\begin{aligned} & 0.939 * * * \\ & (0.107) \end{aligned}$ | 2.502*** |
| Stroke | $\begin{aligned} & 0.432^{* * *} \\ & (0.120) \end{aligned}$ | 1.519*** | $\begin{aligned} & 0.299^{* *} \\ & (0.121) \end{aligned}$ | 1.332** | $\begin{aligned} & 0.572^{* * *} \\ & (0.141) \end{aligned}$ | 1.745*** | $\begin{aligned} & 0.438^{* * *} \\ & (0.124) \end{aligned}$ | 1.522*** |
| Diabetes | $\begin{aligned} & 0.345 * * * \\ & (0.088) \end{aligned}$ | 1.393*** | $\begin{aligned} & 0.530^{* * *} \\ & (0.153) \end{aligned}$ | 1.674*** | $\begin{aligned} & 0.374^{* * *} \\ & (0.100) \end{aligned}$ | 1.433*** | $\begin{aligned} & 0.339 * * * \\ & (0.087) \end{aligned}$ | 1.391*** |
| Lung disease | $\begin{aligned} & 0.247^{* *} \\ & (0.108) \end{aligned}$ | 1.265** | $\begin{aligned} & 0.363^{*} \\ & (0.189) \end{aligned}$ | 1.419* | $\begin{aligned} & 0.283^{* *} \\ & (0.119) \end{aligned}$ | 1.311** | $\begin{aligned} & 0.239 * * \\ & (0.100) \end{aligned}$ | 1.262** |
| Alzheimer | $\begin{aligned} & 0.459 * * * \\ & (0.154) \end{aligned}$ | 1.560*** | $\begin{aligned} & 0.210^{* * *} \\ & (0.051) \end{aligned}$ | 1.219*** | $\begin{aligned} & 0.409^{*} \\ & (0.211) \end{aligned}$ | 1.485* | $\begin{aligned} & 0.476 * * * \\ & (0.156) \end{aligned}$ | 1.567*** |
| Days at hospital (last year) | $\begin{aligned} & 0.012^{* * *} \\ & (0.001) \end{aligned}$ | 1.002*** | $\begin{aligned} & 0.015^{* * *} \\ & (0.002) \end{aligned}$ | 1.005*** | $\begin{aligned} & 0.014^{* * *} \\ & (0.002) \end{aligned}$ | 1.004*** | $\begin{aligned} & 0.014^{* * *} \\ & (0.001) \end{aligned}$ | 1.003*** |
| Days in hospital last year (mental health) | $\begin{aligned} & 0.007^{* * *} \\ & (0.001) \end{aligned}$ | 1.260*** | $\begin{aligned} & 0.008^{* * *} \\ & (0.002) \end{aligned}$ | 1.264*** | $\begin{aligned} & 0.007^{* * *} \\ & (0.001) \end{aligned}$ | 1.258*** | $\begin{aligned} & 0.008^{* * *} \\ & (0.002) \end{aligned}$ | 1.262*** |
| Days in other institutions (last year) | $\begin{aligned} & 0.011^{* * *} \\ & (0.002) \end{aligned}$ | 1.001*** | $\begin{aligned} & 0.011^{* * *} \\ & (0.002) \end{aligned}$ | 1.004*** | $\begin{aligned} & 0.011^{* * *} \\ & (0.002) \end{aligned}$ | 1.007*** | $\begin{aligned} & 0.012^{* * *} \\ & (0.002) \end{aligned}$ | 1.009*** |
| Visits to general practitioner (last year) | $\begin{aligned} & -0.002 \\ & (0.006) \end{aligned}$ | 0.988 | $\begin{aligned} & -0.002 \\ & (0.007) \end{aligned}$ | 0.988 | $\begin{aligned} & 0.002 \\ & (0.006) \end{aligned}$ | 0.992 | $\begin{aligned} & -0.003 \\ & (0.005) \end{aligned}$ | 0.989 |
| Sedentary lifestyle | $\begin{aligned} & 0.851^{* *} \\ & (0.079) \end{aligned}$ | 2.301*** | $\begin{aligned} & 0.834^{* * *} \\ & (0.143) \end{aligned}$ | $2.264^{* * *}$ | $\begin{aligned} & 0.781^{* * *} \\ & (0.088) \end{aligned}$ | 2.148*** | $\begin{aligned} & 0.858^{* * *} \\ & (0.082) \end{aligned}$ | 2.304*** |
| Smokes now | $\begin{aligned} & 0.420^{* * *} \\ & (0.108) \end{aligned}$ | 1.501*** | $\begin{aligned} & 0.405^{* * *} \\ & (0.167) \end{aligned}$ | 1.465*** | $\begin{aligned} & 0.359^{* * *} \\ & (0.113) \end{aligned}$ | 1.413*** | $\begin{aligned} & 0.429 * * * \\ & (0.110) \end{aligned}$ | 1.504*** |
| Smoked before, not now | $\begin{aligned} & 0.312^{* * *} \\ & (0.107) \end{aligned}$ | 1.347*** | $\begin{aligned} & 0.326 * * \\ & (0.143) \end{aligned}$ | 1.368** | $\begin{aligned} & 0.224^{*} \\ & (0.118) \end{aligned}$ | 1.237* | $\begin{aligned} & 0.314^{* * *} \\ & (0.110) \end{aligned}$ | 1.348*** |
| Drinks alcohol daily, 5-6 times/week | $\begin{aligned} & 0.215^{* *} \\ & (0.081) \end{aligned}$ | 1.219** | $\begin{aligned} & 0.226 \\ & (0.089) \end{aligned}$ | 1.230 | $\begin{aligned} & 0.239 * * \\ & (0.094) \end{aligned}$ | 1.251** | $\begin{aligned} & 0.217^{* *} \\ & (0.083) \end{aligned}$ | 1.220** |
| No education | $\begin{aligned} & 0.424^{* * *} \\ & (0.156) \end{aligned}$ | 1.508*** | $\begin{aligned} & 0.533^{* * *} \\ & (0.179) \end{aligned}$ | 1.681*** | $\begin{aligned} & 0.396^{* *} \\ & (0.165) \end{aligned}$ | 1.467** | $\begin{aligned} & 0.428^{* * *} \\ & (0.158) \end{aligned}$ | 1.509*** |
| Primary education | $\begin{aligned} & 0.452^{* * *} \\ & (0.130) \end{aligned}$ | 1.550*** | $\begin{aligned} & 0.388^{* *} \\ & (0.172) \end{aligned}$ | 1.455** | $\begin{aligned} & 0.425^{* * *} \\ & (0.133) \end{aligned}$ | 1.509*** | $\begin{aligned} & 0.457^{* * *} \\ & (0.132) \end{aligned}$ | 1.552*** |
| Lower secondary education | 0.187 | 1.190 | 0.255 | 1.275 | 0.182 | 1.185 | 0.195 | 1.193 |



Standard errors between parenthesis. All regressions include time fixed effects and country fixed effects and are weighted using calibrated sampling weights.


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[^2]:    ${ }^{1}$ De Bresser, (2017) draws on two Dutch surveys, administered to the same respondents during the same month, documenting similar relationships between socio-demographic covariates and the objective and subjective survival expectations.
    ${ }^{2}$ Some different datasets, such as the Health and Retirement Survey (HRS) in the United States and the Survey of Health Ageing and Retirement in Europe (SHARE) contain a specific module with an end-of-life panel component which exhibits to date a reasonable sample with a reasonable response rate. Some studies have already documented evidence of an age specific distribution of subjective and objective survival in the United States (Bissonnette et al. 2017, Palloni and Novak 2016).

[^3]:    ${ }^{3}$ Alternatively, behavioural biases such as longevity optimism, or tendency to view the future more in one's favour, can underpin individuals' differences in survival expectations.
    ${ }^{4}$ We use the stock sampling approach of Jenkins (1995), which allows us to change the unit of analysis from the individual to the time at risk of death, and in turn simplifies the complex sequence of likelihoods to the standard estimation of a binary outcome.
    ${ }^{5}$ So far, most of the literature examines how survival expectations are modified by changes in health behaviours (Smith et al., 2001), genetic information (Hurd and McGarry, 1995; Perozek, 2008), better knowledge of parents' health (van Doorn and Kasl, 1998) and health-related experiences in life (Benítez-Silva, 2008; Costa-Font and Costa-Font, 2011). Previous evidence supports the idea that subjective survival expectations are good predictors of longevity (Hurd and McGarry, 1995, 2002).

[^4]:    ${ }^{6}$ For instance, Salm (2010) found that a $1 \%$ higher subjective mortality rate was associated with an annual decrease in consumption of non-durable goods of $1.8 \%$. Groneck et al. (2017) claim that underestimation of young age and overestimation of old age survival probabilities might give rise to the occurrence of over and under saving, respectively. Puri and Robinson (2007) found that overestimated survival (from life tables) gives rise to more conservative saving behaviours (more savings). However, they identify a non-linearity at the top share of the survival distribution and do not examine optimism.

[^5]:    ${ }^{7}$ The SHARE target population consists of all persons aged 50 years and older at the time of sampling who have their regular domicile in the respective SHARE country, which contains micro level information on demographics, socio-economic status, health status, social and family networks. Persons are excluded if they are incarcerated, hospitalized, or out of the country during the entire survey period or unable to speak the country's languages (Bergman et al., 2019).

[^6]:    ${ }^{8}$ Calibrated sampling weights are missing for respondents younger than 50 years (i.e., age-ineligible partners of an age eligible respondent) and those with missing information on the set of calibration variables (i.e., age, gender and NUTS1 code).
    ${ }^{9}$ The "End of life" module in the event of death is completed between two waves (questions are answered by a proxy respondent). Using the unique identifier assigned to each respondent, it is possible to link the "End of life" module with the "Main questionnaire" of the wave immediately preceding it.
    ${ }^{10}$ This last step configures our final samples: 2,040 for the deceased's subsample and 67,860 for the survivors' subsample, and we test if there is a sample selection problem due to the exclusion of those who have not reported their survival expectations
    ${ }^{11}$ To cope with potential selection bias generated by non-response and panel attrition, SHARE suggests using calibrated sampling weights using the methodology designed by Deville and Särndal (1992).

[^7]:    ${ }^{12}$ Two reasons may explain this underestimation: (i) the institutionalized population is missing in the first wave of SHARE, and although respondents are followed into institutions in the second wave, the institutionalization rate in SHARE is very low, (ii) except for France, the mortality follow-up is not based on registers, but on the "End-of-life interview".
    ${ }^{13}$ We only reject the null hypothesis when comparing follow-up and non-follow-up respondents between waves 2 and 4 (which could be justified given the greater time interval between both waves).

[^8]:    ${ }^{14}$ This fact may be explained by the combination of a high life expectancy ( 80.2 years in 2015) and being one of the countries in the world with the highest number of centenarian population. Source: Health status - Life expectancy at birth - OECD Data.
    ${ }^{15}$ If we consider the classification of countries proposed by Lewis (1996), we can see that: Group III (Austria and Denmark) exhibits the highest SSP for target ages 75, 80 and 85, whereas Group IV (Italy and Spain) shows the highest SSP for the oldest cohorts. In contrast, Group II (Belgium and France) is the group with the lowest SSP for all targets except $\mathrm{T}=90$ years (see classification of countries on Table A4). Table B5 reports survival expectations by gender across different target ages. Importantly, we find that among the survivor's women tend to exhibit higher subjective survival probabilities, but the effect seems to decline when target ages are expanded, and at age 90 men have higher survival probability among survivors.
    ${ }^{16}$ Gender, age, age squared, level of education, marital status, number of days stayed in hospital during the last year (surgery or medical tests), number of days stayed at hospital during the last year (mental health), number of

[^9]:    days stayed at other institutions during the last year, number of visits to general practitioner during the last year, relation with economic activity, living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in 1,000 PPP).
    ${ }^{17}$ Ikeda et al., (2006) showed that the older the age of death of both mothers and fathers, the lower the probability of death for adult children aged between 40 and 79 years. It also seems that longevity is more strongly associated

[^10]:    with maternal death age than parental death age and that mother's longevity reduces the incidence of some pathologies such as pulmonary disease or hypertension (Goldberg et al. 1996; Gjonca and Zaninotto 2008).
    ${ }^{18}$ The predicted kernel density function (using the imputed values for the living parents' subsample) is displaced to the right as compared to the density function for deceased parents. Imputed age of death is around two years higher as compared to parents' reported age, regardless of the adult's child gender.
    ${ }^{19}$ Including gender, age, age squared, level of education (no education, primary education, lower secondary education, higher secondary education, post-secondary non-tertiary education, and tertiary education (omitted)), marital status (married/cohabiting, separated/divorced, single and widow (omitted)).
    ${ }^{20}$ Finally, a third specification considers a series of controls for economic activity Such as employment retirement status, home working, living in a rural area/village/small town, living alone, wealth and income (adjusted by household size and in 1,000 PPP).
    ${ }^{21}$ The rejection of the null hypothesis for the Durbin-Wu Hausman test indicates that the OLS estimations are inconsistent and that the IV technique is required. The Stock and Yogo critical values at 5\% are lower than the Cragg-Donald statistic, which tests for weak instruments. Finally, as we have one endogenous variable and two

[^11]:    instruments, we perform an overidentification test. The p-value from the Sargan test shows that the instruments are uncorrelated with the error term of the outcome equation, and thus are valid instruments.
    ${ }^{22}$ https://www.gaisma.com/en
    ${ }^{23}$ http://data.un.org/Data.aspx?d=CLINO\&f=ElementCode\%3A15

[^12]:    ${ }^{24}$ The sample of respondents to this question is 69,315 as compared with the total sample which amounts to 69,900 observations.

[^13]:    ${ }^{25}$ Being a smoker was found to be equivalent to being at least four years older in terms of its negative effects on two-year survival (Wang, 2014).
    ${ }^{26}$ The high impact of cancer coincides with previous evidence by Hurd and McGarry (2002) who found that cancer was the strongest predictor of mortality, increasing two-year mortality rate by $150 \%$.
    ${ }^{27}$ The hazard ratio for no education and primary education are $53.1 \%$ and $57.5 \%$ higher than the omitted category (tertiary education). Furthermore, the hazard rate for those who have 1,000 PPP additional units of adjusted wealth is $66.3 \%$ of the hazard rate for those who do not have such wealth.

[^14]:    ${ }^{28}$ For example, for the age cohort 50-64, we estimate an average SSP of reaching target age T=75 years is 50.88 for the deceased subsample and 68.06 for the survivor's subsample. In contrast the objective survival probability of reaching target age $\mathrm{T}=75$ years is 34.64 for the deceased subsample and 56.57 for the survivor's subsample. This implies that the gap between SSP and OSP is higher for the deceased subsample ( 16.24 pp vs. 11.51 pp for the survivors).
    ${ }^{29}$ Similarly, Grevenbrock et al. (2016) finds that cognitive impairments cause upward biases in survival beliefs for the oldest age group (age group 85-89) which leads to an overestimation of survival chances by about 15 percentage points. An alternative explanation proposed by Kahneman and Tversky (1979) is that individuals tend to overweight small probabilities and underweight large probabilities.

[^15]:    ${ }^{30}$ See as Kutlu-Koc and Kalwij (2017) and Teppa and Lafourcade (2013) found this for the Netherlands; O’Donnell et al. (2008) for the UK; and Perozek (2008) and Bissonnette et al. (2017) for the US
    ${ }^{31}$ The first result is in line with previous evidence: (i) Wenglert and Rosén (2000) and Wu et al. (2013) found that women are more pessimistic than males in projecting survival probabilities, (ii) Hurd and McGarry (2002) and Gan et al. (2005) observed that men tended to overestimate their lifetimes.
    ${ }^{32}$ Group I: Italy and Spain. Group II: Belgium and France. Group III: Austria and Denmark. Group IV: Germany, Sweden and Switzerland.
    ${ }^{33}$ Table D3 reports the results of the regressions introducing the four instruments, $X_{\text {itc }}$ (vector of explanatory variables), wave and country fixed effects. In most regressions, the four instruments are not significant, confirming that parents' living status only affects offspring's behaviors through the channel of LO.

[^16]:    ${ }^{34}$ This result may result from the fact that smokers perceive that they are more likely to experience certain diseases (Sloan et al., 2001) and it is consistent with rational addiction theory according to which smokers are forward looking individuals who internalize the detrimental effects of smoking (Becker and Murphy, 1988).
    ${ }^{35}$ These effects entail a $71.1 \%$ ( $94.83 \%$ ) increase (decrease) with respect to the mean probabilities.

[^17]:    ${ }^{36}$ Table B3 reports the estimated coefficients and Table B14 reports the effect of a one standard increase of LO and ME.
    ${ }^{37}$ Table B4 reports the estimated coefficients and Table B15 reports the effect of a one standard increase of LO and ME.

[^18]:    ${ }^{38}$ These effects entail an average change in $45 \%$ for bonds, $76 \%$ for shares, $50 \%$ for mutual funds (specifically, $16.2 \%$ for funds invested mostly in stocks; $-19.80 \%$ for funds equally divided in stocks and bonds; $52.3 \%$ for funds mostly invested in bonds) and 69.2\% for retirement accounts.
    ${ }^{39}$ These effects entail and average decrease by $18.66 \%$ (bank accounts), $37.17 \%$ (stocks), $50.05 \%$ (mortgage) and 46.93\% (other debts) with respect to mean probabilities.
    ${ }^{40} \mathrm{~A}$ one standard deviation changes in $\mathrm{BSE}<0$ reduces investment in bonds by 2.70 pp ( 0.65 pp as compared to BSE $>0$ ), mutual funds by 4.69 pp ( 0.9 pp as compared to BSE>0), total wealth by 11.55 ( $1,000 \mathrm{PPP}$ ), and increases life insurance uptake by 6.12 pp ( 0.86 pp as compared to $\mathrm{BSE}>0$ ).

[^19]:    ${ }^{41}$ Table B5 reports the estimated coefficients and Table B16 reports the effect of a one standard increase of BSE and BME

[^20]:    ${ }^{42}$ In order to correct for focal responses several alternatives have been proposed: (i) bayesian updating mechanism to smooth focal points (Gan et al, 2005), (ii) replacement of focal point answers at zero with 0.01 and focal point answers at 100 with 99.9 (Picone et al., 2004), (iii) instrumental variables for subjective mortality expectations in order to adjust for focal answers (Delavande et al., 2006). However, other authors have decided against correcting probabilities. Khwaja et al. (2007), Salm (2010) and Post and Hanewald (2013) consider that focal responses are still what agents base their decisions on. Hill et al. (2005) consider that an answer of 50 per cent may be the true probability, if the respondent believes that the event of dying is equally likely to occur or not to occur.
    ${ }^{43}$ In this re-estimation process, we again test which hazard function fits better to the data. Cox-Snell residuals and information criteria point to the Weibull function as before (graphs are available upon request; AIC and BIC shown in Table C1).

[^21]:    ${ }^{44}$ With the reduced sample ( $\mathrm{N}=22,198$ ) we have re-estimated the discrete-time hazard model (see Table B1 for a description of the sample). The information criteria (reported in Table C1) and the Cox-Snell residuals (available upon request) show that the Weibull specification provides the best fit. The estimated coefficients and hazard ratios for the preferred specification are reported in Table B8 ( $4^{\text {th }}$ column).
    ${ }^{45}$ SHARELIFE collects respondent's life histories, including medical tests, and the disadvantage is that merging SHARELIFE with our sample implies an important reduction in the number of observations (from 69,900 to $33,269)$.
    ${ }^{46}$ The rationale refers to a link between parent's history of health conditions, such as blood pressure and diabetes, and the probability that children suffer the same pathology (Hakonarson et al. 2007; Newtwon-Cheh et al., 2009). Therefore, offspring who have a higher risk of suffering a heart attack or developing diabetes may be more concerned about their survival probabilities.

[^22]:    ${ }^{47}$ Consistently, Table B8 ( $5^{\text {th }}$ column) reports the results for the discrete-time hazard model using the new sample ( $\mathrm{N}=33,269$ ).

[^23]:    ${ }^{48}$ The first column corresponds to the estimates using the entire sample and current instruments. The second and third columns correspond to estimations of those individuals for whom lagged observations are available: using current instruments in the second column and delayed instruments in the third column.
    ${ }^{49}$ In the right column figures confidence intervals are presented. The x axis measures how strong the violation of the exclusion restriction needs to be for the instrument to turn insignificant. In all figures, the confidence intervals do not include the value 0 (red line), so we can infer that the IV estimations are robust to possible violations of the exclusion restriction

[^24]:    ${ }^{50}$ In some cases, the effect of a one standard deviation increase of negative BSE is higher as compared to that of positive BSE. For example, an increase of negative BSE generates a strong disinvestment in bonds, mutual funds (mostly in bonds) and a significant decrease in total wealth.

[^25]:    ${ }^{51}$ Self-reported health is a good indicator of morbidity and a powerful predictor of future health and mortality (Burström and Fredlund, 2001; Van Doorslaer and Gerdtham (2003).

[^26]:    ${ }^{\text {c }}$ Resulting from merging two consecutive final sample datasets obtained in the previous table.
    ${ }^{\mathrm{d}}$ Missing, do not know, refusal.
    Survivors sample: alive in the second wave; Deceased sample: died between both waves; Not follow-up respondents: only appear in one wave.

[^27]:    Wilcoxon Rank sum test. Null hypothesis: both samples are from are from populations with the same distribution.

[^28]:    Standard errors between parenthesis. All regressions include time fixed effects and country fixed effects and are weighted using calibrated sampling weights.

[^29]:    52 Age, gender, level of education, body mass index, chronic illness, smoking and drinking habits, feeling depressed, days stayed at hospital during last year for surgery or medical tests, days stayed at hospital during last year due to mental health problems, number of days stayed in other institutions other than a hospital or a nursing home during last year (i.e., institutions for rehabilitation, convalescence), visits to general practitioner during last year, sedentary lifestyle, living alone, living in a rural area, being employed, adjusted income and wealth, country and time fixed effects. The definition of each one is shown on Table A1.

[^30]:    ${ }^{53}$ These figures have been obtained using the command plausexog proposed by Clarke (2014) for STATA.

