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## How Much Do Households Dislike Local Density? And Do Developers Fully Consider Their Preferences? Evidence from a Policy Change in Singapore

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## ABSTRACT

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# How Much Do Households Dislike Local Density? And Do Developers Fully Consider Their Preferences? Evidence from a Policy Change in Singapore

This paper measures how much households dislike density in their immediate surroundings. Using transaction and administrative data in Singapore, and exploiting the introduction of a regulation that restricted the number of housing units for certain land lots, we find that households do indeed discount density: a 10% increase in within-development density decreases price per square meter by up to 4%. Further, we find that the mean price per square meter of the average development increased by 1 to 3% after the regulation was introduced, while the amount of built-up space remained constant. The increase in total revenue suggests that developers may underestimate the externality caused by density. Our results are particularly relevant during the lockdowns and social distancing of the coronavirus pandemic.

**JEL Classification:** R20, R21, R38

**Keywords:** density, regulation, land-use policy, externalities

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# 1 Introduction

As the world continues to urbanize, cities have become more dense, leading to externalities such as traffic congestion, air pollution, and loss of open spaces. In response, land-use regulations have become widespread as cities attempt to control the intensity of residential development and strike a balance between fully utilizing land and ensuring livability.

An extensive literature has explored land-use restrictions (Gyourko and Molloy, 2015; Molloy, 2020), finding that such regulation leads to a smaller stock of housing and higher prices in the United States (Glaeser and Gyourko, 2018; Zabel and Dalton, 2011; Ihlanfeldt, 2007) and internationally (Hilber and Vermeulen, 2016; Tan et al., 2020). There are fewer papers that focus on the economic effects of density itself, with Ahlfeldt and Pietrostefani (2019), Ahlfeldt et al. (2015), Ciccone and Hall (1996), and Brueckner and Largey (2008) being some prominent examples, and even fewer papers that consider how *local* density, that is, the density of the immediate surroundings, affects households, which is the focus of this paper.<sup>1</sup>

Local density is particularly relevant during the coronavirus pandemic, since it has highlighted the value of space as we work and spend much of our time at and near our homes during lockdowns and social distancing. Indeed, early speculation that denser cities would experience more infections has been refined to emphasize the importance of crowdedness of the immediate environment, for instance, at home, in dormitories, in offices, and on public transportation (Fang and Wahba, 2020; Zhong and Teirlinck, 2020). The pandemic has also underscored inequalities in living conditions as poorer and more crowded neighborhoods often suffer more than predominantly wealthier and less populated ones (Arango, 2021), and well-to-do city dwellers are more likely to move to suburban or rural locations (Lerner, 2020). Neuroscientists and psychologists even propose that people are learning new emotional experiences during the pandemic: unease around strangers and aversion to crowds (Kiefer, 2020).

In addition to the risk of infectious disease and its associated effects in the new pandemic world, drawing on findings in psychology and sociology we speculate that households dislike local density because of a range of other factors including competition for facilities and common areas,

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<sup>1</sup>Tang and Yiu (2010) and Lee (2016) concluded that buyers pay more when there is more space and less density locally but did not address the endogeneity of density.

lack of privacy, noise, poor sleep, and adverse social interactions that cause stress, conflict, and pathological behavior, which have been found to be more common in crowded areas (Gray, 2001; Gómez-Jacinto and Hombrados-Mendieta, 2002; Chan, 1999).

To summarize succinctly, local density has a deep, continual impact on our physical and mental well-being but there is little economic research that has attempted to measure people’s disutility to it. One sensible approach is to measure how much less housing costs in more dense environments, the approach we take in our research. Providing policymakers with an accurate measure of the discount people place on dense living environments will allow more efficient planning and zoning policies, particularly if policymakers are less aware of the impact of local density compared to other more well-known externalities.

To measure how much households value less density, we constructed a novel dataset from planning documents of relatively low-density residential developments in Singapore and combined it with transaction data.<sup>2</sup> We find that households do place a premium on less crowded environments: a 10% increase in density reduces the price per square meter by up to 4%. To the best of our knowledge our findings are the first to measure the causal effect of local density on prices in relatively low-density developments.<sup>3</sup>

Our measure of density is the number of housing units per meter of lot size, which one would expect to be endogenous in a price regression. For instance, density would be endogenous if developers build more intensely in especially attractive areas. To address such endogeneity we make use of a first-of-its-kind regulation in Singapore introduced in November 2011 in response to developers building an increasing number of very small housing units, also known as “shoebox” units.<sup>4</sup>

The regulation imposed a minimum average size of housing units in developments built on lots with a 1.4 Floor Area Ratio (FAR) constraint, the lowest FAR value of residential land lots in Singapore.<sup>5</sup> By constraining the mean unit size, the regulation decreased the number of dwelling

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<sup>2</sup>The regulation we use to identify the effect of density occurred in pre-pandemic times. A natural extension of our research is to measure whether people’s dislike for density has increased from 2020.

<sup>3</sup>Fesselmeyer et al. (2018) studied the price effects of local density in a mix of low and high density and low and high rise developments using a different identification approach based on Singapore’s geographic mixing of low- and high-density regulated land lots in certain areas.

<sup>4</sup>The Urban Redevelopment Authority (URA) defines shoebox units as units 50 square meters or smaller in size.

<sup>5</sup>The local term for Floor Area Ratio is Gross Plot Ratio (GPR). We use Floor Area Ratio throughout the text

units that could be built on a given 1.4 FAR land lot. Lots regulated with a larger FAR were not effected. Therefore, we use 1.6 to 2.1 FAR developments as a control group so that the interaction between the FAR value and the introduction of the regulation creates a difference-in-differences instrument for density. We explain the regulatory change and the effect on density in detail below, showing that the instrument explains the decreased post-regulation density of 1.4 FAR developments absolutely and also relative to 1.6-2.1 FAR developments.

In addition to estimating the effect of density on price, we consider how developers changed the mix of housing units within a development in response to the density restriction, finding that developers reduced the proportion of shoebox units, while increasing the proportion of medium and large units, in order to meet the minimum average size constraint. The changes in design affected price: the proportion of shoebox units is inversely related to price, while the share of large units is positively related to price.

Finally, we perform a profit analysis for an average 1.4 FAR development, before and after the policy change. Surprisingly, the regulation increased profit even though developers built the same amount of floor area for a given lot size. The reduction in density increased the price per square meter for all sized units, which more than compensated for the decrease in revenue caused by the declining share of small units, which tend to have a higher price per square meter. This begs the question of why did the developers overbuild small units before the regulation. We speculate that a possible reason is that developers were unaware or under-estimated how much households dislike within-development density.

## 2 Institutional Setting and Policy Change

Singapore is a densely populated city-state of 5.68 million people living on a land area of 725.5 square kilometers (280.1 square miles). The vast majority of residents own their own homes, with a high homeownership rate of 90.4% in 2019 (Department of Statistics Singapore, 2020). The housing market is characterized by the co-existence of public housing, apartments in multi-unit buildings that are built by the Housing & Development Board (HDB) and restricted to Singaporean

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to minimize any confusion.

citizens and permanent residents, and private housing in multi-unit buildings that are built by private developers and open to everyone. Locally, these privately developed units are referred to as condominiums. They are located in “developments,” a collection of adjacent buildings, each with many condominiums that share a land lot, name, and facilities. About 80% of the population live in HDB dwellings. Most of the remaining population live in condominiums. We focus on the private residential sector because it is not restricted by residency status or in any other significant way, and it is not subsidized.

Given the land constraints in Singapore, land use and long-term planning are fundamental. The agency in charge of planning, the Urban Redevelopment Authority (URA), uses the Concept Plan to set out the long-term strategic plans that address broader issues for the next 50 years, and uses the more-frequently revised Master Plan to guide development in the medium term and to carry out the goals laid out in the Concept Plan, including setting regulations on density, building form and height, and open space for every land parcel in the country.

As the economy has prospered with a 7% annual growth rate since independence in 1965 and the population has grown from 1.89 million in 1965 to 5.69 million in 2020, Singapore has attempted to balance growth and the negative effects of “too much” density. The primary regulatory tool to restrict density has been the Floor Area Ratio, the maximum amount of floor space per land space allowable on a given plot of land.

Broadly speaking, the density of residential developments can be categorized into two groups based on the allowable FAR. Low-density developments are relatively low-rise multifamily housing built on lots with a FAR constraint of less than 2.8, with the most common FAR values being 1.4 and 2.1. High-density developments are typically high-rise multifamily housing with FAR values of 2.8 or higher. The policy we use in our identification strategy impacts 1.4 FAR developments, and, as such, we focus exclusively on low-density housing in this paper (in contrast to Fesselmeier et al. (2018), which included both low-density and high-density housing in their analysis of density effects).

In recent years, the URA has introduced restrictions on the number of housing units per lot as an additional regulatory tool. The first such restriction was announced by the URA on November 23, 2011 with effect on the following day (Urban Redevelopment Authority, 2011). The maximum



allowable number of housing units for most 1.4 FAR lots was determined by the following formula:

$$\bar{X} \leq \frac{FAR \times SA}{70} \quad (1)$$

where  $\bar{X}$  denotes the number of housing units,  $FAR$  is the Floor Area Ratio, and  $SA$  is the lot size in square meters. The denominator of 70 square meters is the minimum average unit size allowable in the development. Prior to this policy, there were no restrictions on the number of housing units or the average unit size.

In addition, the URA identified certain neighborhoods of 1.4 FAR lots for a more restrictive version of the regulation in which average size of units would be 100 sqm (so that the denominator of equation (1) is 100). We refer to these neighborhoods as “special neighborhoods” below.

In 4 November 2012, the URA expanded the housing unit restriction of equation (1) to all new developments outside the downtown area (referred to officially as the Central Area) (Urban Redevelopment Authority, 2012). Given that our focus is on the effects of density in low-density developments, we restrict our sample to developments of 2.1 FAR or lower, with construction plans submitted before November 2012.

### 3 Data

Our empirical analysis uses development characteristics data combined with transaction data. We describe each below.

#### 3.1 Development data

We assembled a dataset of private residential low-density developments (on 1.4 to 2.1 FAR lots) that submitted construction plans to URA from 2010 to November 2012, using URA’s Real Estate Information System (REALIS) database, URA SPACE, an integrated map portal that contains master plan information and planning documents, and OneMap, a comprehensive map of Singapore maintained by the Singapore Land Authority.

Our sample consists of 140 developments: 105 have a FAR of 1.4 and 35 have a FAR between 1.6 and 2.1. Figure 1 plots the developments on a map with postal district boundaries, a geo-

graphic demarcation commonly used by developers and agents. The blue squares are 1.4 FAR developments, and the red triangles are 1.6 to 2.1 FAR developments. We can see that both types of developments are spread out across Singapore. In some robustness checks, we restrict the sample to developments in postal districts with both FAR categories so as to minimize differences in unobserved neighborhood amenities.

We collected the following variables from planning documents from URA SPACE for each development: (1) date of submission of the construction plan, (2) total number of housing units, (3) floor area, and (4) realized floor area ratio. We compute lot size from the ratio of floor area and the realized floor area ratio.

We define the *density* of a development as the total number of housing units in the development per square meter of land. In Panel A of Table 1, we see that the average density of sample developments was 0.023 housing units per square meter of land. Density was higher in the 1.6-2.1 FAR developments (0.026 units per meter of land) than 1.4 FAR developments (0.021 units per meter of land).

Figure 2 plots the densities of 1.4 FAR developments in blue squares and 1.6-2.1 FAR developments in red circles by the date of submission of construction plans. The effective date of the policy change is represented by the dashed vertical line. The blue horizontal lines are the mean densities of 1.4 FAR developments, before and after the policy change. The red horizontal lines are the mean densities of 1.6-2.1 FAR developments, before and after the policy change.

We see that not only did the density of 1.4 FAR developments fall substantially after the policy change but that the fall was more pronounced compared to the increase in density of 1.6-2.1 FAR developments. In other words, the difference in the mean density of 1.4 FAR and 1.6-2.1 FAR developments increased after the policy change, which supports our identification scheme that uses the decrease in 1.4 FAR density after the policy change relative to the unregulated control group of 1.6-2.1 FAR developments.

Further support for the identification scheme is given in Table 2, which computes how the difference in average density by FAR changed after the policy date. In Panel A, we see in the last row of column 1 that the 1.4 FAR developments were 0.135 log points less dense (12.6%) than 1.6 to 2.1 FAR developments. After the policy change, in column 2, the difference increases to 0.438

log points, or 35.5%. In the last row of column 3, the difference in these differences is 0.304 log points, or a 26.2% difference in density, which is significant at 1%.

In Panel B, we make a similar calculation but use a placebo policy change dated one year earlier than the actual policy date. The before period is from the beginning of the sample, January 1, 2010, to the placebo policy date of November 24, 2010. The after period is from that date until the actual policy date, November 24, 2011. In the last row, the differences in density are smaller using the placebo policy date, including the difference-in-differences estimate of -0.129 log points in the last row of column 6 (12.1% difference), which is not significant at any conventional significance level. This evidence suggests that the reduction in density among 1.4 FAR developments after the policy change is driven by the policy change itself.

We collected facilities information such as the presence of a swimming pool, barbecue area, gym, tennis court, sauna, jacuzzi, etc. from condo websites and real estate websites such as propertyguru.com.sg. We collected Construction Quality Assessment System (CONQUAS) data from the Building and Construction Authority, which includes a quality assessment of the structural, architectural, and mechanical engineering works of developments. We computed distance to the nearest Mass Rapid Transit (MRT) station, Singapore's public rail system, and the distance to the Central Business District from geocoordinates extracted from OneMap. Finally, we note that, as a former British colony, Singapore follows the British leasehold system in which land is either freehold, i.e., owned in perpetuity, or leased from a freeholder, usually the government, for a certain number of years. About 23.6% of the developments are on land with 99-year leases. The remaining leases are either 999 years or freehold. For further details on the leasehold system in Singapore, readers may refer to Fesselmeyer et al. (2021).

Summary statistics of these development-level variables for the sample and broken down by FAR can be found in Panel A of Table 1.

### **3.2 Transaction data**

Transaction data is available from the REALIS database, which contains records of caveats of private residential properties. The data includes housing-unit level information such as the date of sale, transaction price, floor area, age, and floor. We collected new transactions (transactions

between the developer and the first buyer) for the 140 developments included in the analysis. The 20,713 transactions span from 2010 to 2017. Because it takes several years to sell all units in a development, we observe transactions of both effected and uneffected developments for several years after the submission of construction plans.

Pabel B of Table 1 contains the summary statistics of housing unit-level variables, including a breakdown by FAR. The main outcome variable of interest is transaction price per square meter (deflated to 2000 SGD values). The average is SGD 9,626.65, or about USD 6,876. On average, condominiums in 1.4 FAR developments sold for a higher price than those in 1.6-2.1 FAR developments, SGD 9,777.40 per square meter vs. SGD 9,475.49 per square meter (USD 6,983 vs. USD 6,768). The average deflated transaction price was SGD 804,677.80 (USD 574,770), with 1.4 FAR condominiums selling for SGD 772,926.70 and 1.6-2.1 FAR condominiums sold for SGD 836,511.80 (USD 552,091 and USD 597,508). The differences in density and unit size could at least partially explain the difference in prices. Units in 1.6-2.1 FAR developments tend to be larger. They averaged around 90 square meters in size compared to 82 square meters for the 1.4 FAR condominiums. We also note that the 1.4 FAR buildings are shorter than 1.6-2.1 FAR buildings, as expected. The average floor of 1.4 FAR condominiums is 3.3 while the average floor of 1.6-2.1 FAR condominiums is 8.4.

## 4 Empirical Strategy

One goal of this paper is to estimate the impact of density on residential prices. The main identification problem is that density is likely endogenous because any unobservable, e.g., neighborhood quality, that affects price is also likely to affect how many units the developer decides to build. In order to estimate the causal effect of density on price we use the policy restriction on the number of housing units in 1.4 FAR developments to construct an instrumental variable, measuring how density changed after the policy for 1.4 FAR developments compared to the control group of 1.6-2.1 FAR developments that were unaffected by the policy change, while including a large set of unit- and development-level characteristics and geographic and temporal fixed effects.

Specifically, the main equation of interest is:

$$\ln P_{ijt} = \beta_0 + \beta_1 FAR14_j + \beta_2 POST_t + \delta \ln D_j + \gamma X_{ij} + \tau_{jt} + \epsilon_{ijt}, \quad (2)$$

where  $P_{ijt}$  is the (deflated) price per square meter of condominium  $i$  in development  $j$  in period  $t$ ,  $FAR14_j$  is a 1.4 FAR development dummy variable,  $POST_t$  is a post-policy change dummy variable,  $D_j$  is the density of the development (the number of housing units in the development divided by the lot size in meters),  $X_{ij}$  are observed characteristics,  $\tau_{jt}$  are time and location fixed effects, and  $\epsilon_{ijt}$  is an idiosyncratic error. The parameter of interest is  $\delta$ , the price elasticity with respect to density.

The first stage is given by

$$\ln D_j = \alpha_0 + \alpha_1 FAR14_j + \alpha_2 POST_t + \alpha_3 FAR14_j \times POST_t + \mu X_{ij} + \tau_{jt} + \nu_{ijt}, \quad (3)$$

where  $\nu_{ijt}$  is an idiosyncratic error. The interaction term  $FAR14_j \times POST_t$  is our instrument for density, allowing us to infer the exogenous variation in  $\ln D_j$ . Given what we know of the Singapore housing market, including what we saw in Figure 2 and Table 2, and the intention of the regulation, the expectation is that  $\alpha_1 < 0$  and  $\alpha_3 < 0$ , that is, 1.4 FAR developments were less dense than 1.6 to 2.1 FAR developments before the policy change and that the difference in densities between the two FAR categories grew larger after the policy change.

## 5 Results

### 5.1 Primary regression results and robustness checks

Table 3 reports our primary results. Columns 1 and 2 include the estimated effect of log density on log price per square meter, including housing-unit characteristics (floor fixed effects and a quadratic of area in meters), sales year fixed effects, and postal district fixed effects. Columns 3 and 4 add development-level characteristics (a 99-year leasehold dummy, a special-neighborhood dummy, the log distance to the CBD and to the nearest MRT station, the number of facilities in the development, and whether the development was CONQUAS assessed and an interaction with

CONQUAS quality score). Columns 5 and 6 add a more flexible time trend, with sales year fixed effects for each postal district. The reason for adding these additional explanatory variables is to control for the potential correlation between housing quality and development density. The odd columns 1, 3, and 5 contain OLS estimates. The even columns 2, 4, and 6 contain two-stage least squares estimates. The third-to-last row of the table reports the  $F$ -statistic of the instrument in the first stage, a check on whether the instrument has “enough” explanatory power.

The specification in column 4 with housing-unit and dwelling-unit characteristics and a common time trend of sales year fixed effects, estimated using two-stage least squares, is our preferred approach. We use it and some variants in several robustness checks and related regressions below.

There is a clear pattern across the different results in Table 3: increased density caused price to be lower. The coefficient estimate on log density is negative and significant for all two-stage least squares estimates and for 2 out of the 3 OLS estimates. Consider column 2, with an estimate of -0.815 on log density, which is significant at the 1% level, indicating that a 1% increase in density leads to 0.56% decrease in price per square meter. The results reported in columns 4 and 6 show that controlling for various observed housing characteristics indeed reduces the estimated negative impact of density on housing price, which is consistent with the hypothesis that developers tend to build higher quality houses in low density developments. In the specification with the most controls, column 6, the density elasticity is -0.481. We also note that the  $F$ -statistic on the instrument is around 17 or higher, indicating that the instrument is not weak according to the 10 rule-of-thumb (Stock and Yogo, 2003), and that the two-stage least square estimate on log density is more negative than the corresponding OLS estimate for each specification pair, which is consistent with the expectation that OLS is upward biased due to unobserved quality.

In Table 4, we re-estimate the regressions from Table 3 but restrict the sample to transactions of developments in postal districts that contained at least one 1.4 FAR development and one 1.6-2.1 FAR development. This robustness check is meant to ensure that each treated development has at least one corresponding control development nearby to minimize the influence of unobserved neighborhood amenities. The results are similar: density has a negative effect on price. Further, the two-stage least square estimates in columns 2, 4, and 6 are very similar in magnitude and significance to those in Table 3.

In Table 5, we examine whether the impact of density on price depends on the size of the housing unit purchased. In column 1, we re-estimate the main specification on a sample consisting of only shoebox units. We see that an increase in density of 1% decreases price per square meter by 0.6%. In column 2, we estimate the model using a subsample of units larger than 50 square meters (that is, we exclude shoebox units). The estimate is similar to the Table 3, column 4 estimate, indicating a decrease in price per square meter of about 0.3%. The larger effect estimated in the shoebox subsample might reflect the premium shoebox dwellers with a small amount of private space put on having fewer neighbors with whom to share the common areas of the development.

## 5.2 Changes in development composition

In this subsection, we consider how developers adjusted their designs, beyond density, in order to meet the new regulation. Specifically we consider whether developers changed the composition of the developments they built after the policy change by examining the distribution of shoebox units (50 square meters or less), medium-sized units (more than 50 square meters to 106 square meters, the 75<sup>th</sup> percentile), and large units (106 square meters and larger) across FAR category and time. We then measure how composition changes affected price.

Figures 3 to 5 contain histograms of the composition of developments by FAR and period. Panel (a) of Figure 3 contains the share of shoebox units in 1.4 FAR developments before the policy. There was a wide range of outcomes. The proportion of shoebox units in some developments was very small but there were many developments in which over 50% of the housing units, some even higher than 90%, which were shoeboxes. We can contrast this with the distribution of shoeboxes in 1.6-2.1 FAR developments before the policy: a large majority of developments had very small proportions of shoebox units.

Panel (b) contains the proportion of shoebox units in 1.4 FAR developments after the policy change. We see, compared to panel (a), that there was a significant decrease in the share of developments with a substantial proportion of shoebox units. Panels (a) and (b) clearly show that developers responded to the policy by building fewer very small units. Panel (d) shows that the developers of 1.6-2.1 FAR developments did not change their layouts substantially.

Figure 4 contain similar histograms of medium-sized units. Compared to the shoebox his-

tograms, there is a less substantial difference between 1.4 FAR developments and 1.6-2.1 FAR developments before the policy change. Panels (a) and (b) show that there was a decrease in the share of 1.4 FAR developments with a small proportion of medium-sized units after the policy change.

Figure 5 contains histograms of large units. These histograms show the least heterogeneity. The 1.4 FAR developments and the 1.6-2.1 FAR developments show a similar distribution before and after the policy change. Both development types exhibit an increase in the proportion of large units after the policy change.

Overall, we can conclude from the histograms that developers of 1.4 FAR lots were building developments with a very high proportion of shoeboxes before the introduction of the policy. In order to adhere to the regulatory requirement, developers decreased the proportion of shoebox units, while also increasing the proportion of medium-sized and large units, thereby increasing the average unit size of a development.

In Table 6, we consider how unit composition within developments affected price. In column 1 our explanatory variable of interest is the percentage of shoebox units in the development, column 2 the percentage of medium-sized units in the development, and column 3 the percentage of large units in the development. We continue to use the interaction between the 1.4 FAR dummy and the post-policy change dummy as an instrument.

The results are intuitively appealing, given our findings in the density regressions and the histograms of size distributions. In column 1, we see that a 1 percentage-point increase in shoebox units decreased price per square meter by 0.3%, and the result is statistically significant. In column 2, the proportion of medium-sized units had no significant effect on price (and that the instrument does not provide a strong response, with an  $F$ -statistic of 2.76). In column 3, we see that a greater proportion of large units increased price: a 1 percentage-point increase caused price per square meter to increase by 0.2%.



## 6 Profit analysis

The policy reduced the number of units a developer can build on a 1.4 FAR lot by increasing the required average unit size, which we have seen was achieved by reducing the proportion of small units and increasing the proportion of medium-sized and large units. The proportion of shoebox units in 1.4 FAR developments declined by 14.7 percentage points from 32.4% in the pre-policy period to 17.7% in the post-policy period, and the proportion of medium-sized and large units increased by 8.2 and 6.2 percentage points, respectively. Because price per square meter is decreasing in unit size, that is, price per square meter is higher for smaller units than larger units, the change in size composition away from small units could have reduced the mean price per square meter of the development, which could explain why developers only reduced density after the regulation was introduced. Alternatively, a reduction in density could have instead benefited developers because, all else equal, it raises the price for units of all sizes because households dislike density. If the developers or their sales staff did not fully internalize this externality because they were not aware of its extent or were not incentivized to do so it is possible that the regulation actually made developers better off.

To examine this possibility more closely, we predict the price of a 1.4 FAR development with the average size composition before and after the policy change, finding that the price per square meter for the entire development was higher after the implementation of the regulation. Moreover, built-up area did not change. The average realized gross plot ratio of 1.4 FAR developments was constant at 1.51 (which includes bonus allowances that developers took advantage of, before and after the policy). Since built-up area did not change, cost differences were minimal, and the predicted increase in price per square meter indicates that profit per square meter increased. This suggests that developers built too many shoebox units in the pre-policy period, foregoing some profit.

We present the calculations in Tables 7 and 8. In Table 7, we categorize units into 3 sizes, shoebox, medium-sized, and large units. In Table 8, we use a finer categorization of unit sizes, creating 10 bins based on floor area percentiles of the transaction data.

Column 1 contains the mean area in square meters of each category. Columns 2 and 3 contain

the average proportion of each size category, before and after the regulation, and columns 4 and 5 contain the mean log density of 1.4 FAR developments, before and after the regulation. Columns 6 and 7 contain the predicted log price per square meter using the 2SLS regression estimates in Table 3, column 4 (panel A) and in Table 3, column 6 (panel B) for units with column 1 area and log density reported in columns 4 and 5, respectively. We used median values for other covariates. In columns 8 and 9, we transform the predicted log price in columns 6 and 7 to price per square meter using the formula  $\hat{p} = \exp(\hat{\sigma}^2/2) \exp(\widehat{\ln p})$ , where  $\sigma^2$  is the variance of the error term. Average price in the last row of columns 8 and 9 in each panel is the weighted average price per square meter using price in columns 8 and 9 and the shares in columns 2 and 3.

In both panels of Table 7, we see that price per per square meter is higher after the policy regulation than before. In Panel A, price per per square meter is SGD 7,671.47 vs. SGD 7,587.13, a 1.1% difference, and in Panel B, it is SGD 7,450.64 vs. SGD 7,223.70, a 3.1% difference. The small differences are not surprising otherwise one might expect the developer to be more aware of the externality and/or more active in minimizing its effect.

In Table 8, we see again that the average developer did better under the regulation. Price is 1.1% higher, SGD 7,690.35 in the post-policy change period vs. SGD 7,609.79 in the pre-policy period.

## 7 Conclusion

In this paper, we use the introduction of a regulation governing the average unit size of low-density residential developments in Singapore to measure the causal impact of very local density on housing prices. We find that price per square meter is lower when density is greater, confirming that households dislike within-development density. We also find that developers reduced the proportion of small shoebox units after the policy change both absolutely and in comparison to a control group of developments unaffected by the regulation. Finally, we show that developers of restricted lots received a *higher* overall price per square meter after the introduction of the regulation, and since built up area and construction costs did not change, this indicates that profits also increased. We speculate that this unexpected outcome may have occurred because developers did not understand

the extent of the density externality, which also suggests that policymakers might not either.

The regulation we use to identify the effect of density occurred in pre-pandemic times. A potential extension of our research is to measure whether people's dislike for density has increased since 2020 as lockdowns and social distancing has become part of our lives. We leave this question for future work.

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Table 1: Summary statistics

| Panel I: Development-level summary statistics |                  |         |         |         |               |         |
|---|------------------|---------|---------|---------|---------------|---------|
|   | All developments |         | 1.4 FAR |         | 1.6 - 2.1 FAR |         |
|   | Mean             | Std Dev | Mean    | Std Dev | Mean          | Std Dev |
| Density (housing units per m <sup>2</sup> )   | 0.023            | 0.007   | 0.021   | 0.006   | 0.026         | 0.007   |
| Housing units                                 | 160.1            | 202.5   | 106.9   | 151.8   | 319.7         | 249.3   |
| 99-year leasehold                             | 0.236            | 0.426   | 0.162   | 0.370   | 0.457         | 0.511   |
| Special neighborhood                          | 0.421            | 0.496   | 0.552   | 0.500   | 0.029         | 0.169   |
| Gym   | 0.686            | 0.466   | 0.629   | 0.486   | 0.857         | 0.355   |
| Swimming pool                                 | 0.971            | 0.167   | 0.962   | 0.192   | 1.000         | 0.000   |
| Tennis court                                  | 0.214            | 0.412   | 0.124   | 0.331   | 0.486         | 0.507   |
| Jacuzzi                                       | 0.336            | 0.474   | 0.333   | 0.474   | 0.343         | 0.482   |
| Sauna   | 0.186            | 0.390   | 0.152   | 0.361   | 0.286         | 0.458   |
| Barbecue                                      | 0.807            | 0.396   | 0.790   | 0.409   | 0.857         | 0.355   |
| Total number of facilities                    | 3.200            | 1.405   | 2.990   | 1.411   | 3.829         | 1.200   |
| CONQUAS-assessed                              | 0.393            | 0.490   | 0.267   | 0.444   | 0.771         | 0.426   |
| CONQUAS score                                 | 90.06            | 6.30    | 87.96   | 7.41    | 92.32         | 3.83    |
| Distance to the CBD (m.)                      | 6,816.3          | 3,561.6 | 6,527.3 | 3,098.3 | 7,683.6       | 46,32.9 |
| Distance to MRT station (m.)                  | 809.1            | 406.2   | 812.9   | 383.7   | 797.6         | 473.3   |
| Proportion of shoebox units                   | 0.309            | 0.294   | 0.344   | 0.301   | 0.205         | 0.248   |
| Proportion of medium-sized units              | 0.500            | 0.229   | 0.478   | 0.227   | 0.568         | 0.226   |
| Proportion of large units                     | 0.187            | 0.198   | 0.174   | 0.195   | 0.226         | 0.203   |
| Developments                                  | 140              |         | 105     |         | 35            |         |

| Panel II: Housing unit-level summary statistics |                  |            |            |            |               |            |
|---|------------------|------------|------------|------------|---------------|------------|
|   | All developments |            | 1.4 FAR    |            | 1.6 - 2.1 FAR |            |
|   | Mean             | Std Dev    | Mean       | Std Dev    | Mean          | Std Dev    |
| Transaction price (2000 SGD)                    | 804,677.80       | 451,094.30 | 772,926.70 | 470,915.80 | 836,511.80    | 427,969.60 |
| Price per m <sup>2</sup> (2000 SGD)             | 9,626.65         | 2,769.77   | 9,777.40   | 2,701.62   | 9,475.49      | 2,828.52   |
| Floor   | 5.86             | 4.54       | 3.30       | 1.58       | 8.42          | 5.06       |
| Area (m <sup>2</sup> )                          | 85.91            | 39.07      | 81.76      | 42.76      | 90.08         | 34.48      |
| Sale year                                       | 2012             | 1.14       | 2012       | 1.03       | 2012          | 1.23       |
| Observations                                    | 20,713           |            | 10,370     |            | 10,343        |            |

Table 2: Mean log density by FAR before and after the policy change and a placebo test

|                | Panel A. Actual policy |                   |                   | Panel B. Placebo test |                   |                   |
|----------------|------------------------|-------------------|-------------------|-----------------------|-------------------|-------------------|
|                | Before                 | After             | (2)-(1)           | Before                | After             | (5)-(4)           |
|                | (1)                    | (2)               | (3)               | (4)                   | (5)               | (6)               |
| 1.4 FAR        | -3.850<br>(0.036)      | -4.047<br>(0.008) | -0.197<br>(0.053) | -3.847<br>(.061)      | -3.852<br>(0.009) | -0.006<br>(0.074) |
| 1.6 to 2.1 FAR | -3.715<br>(0.004)      | -3.608<br>(0.008) | 0.107<br>(0.084)  | -3.777<br>(0.010)     | -3.653<br>(0.009) | 0.124<br>(0.143)  |
| Difference     | -0.135<br>(0.078)      | -0.438<br>(0.059) | -0.304<br>(0.098) | -0.070<br>(0.130)     | -0.199<br>(.089)  | -0.129<br>(0.157) |

Note: Each cell contains mean log density, with the standard error in parentheses below. In panel A, the before period includes all developments with plans submitted before the policy date, November 24, 2011. The after period includes all developments with plans submitted after the policy date. In panel B, the placebo test imposes a placebo policy date of November 24, 2010. The before period includes all developments with plans submitted before the placebo policy date. The after period includes all developments with plans submitted after the placebo policy date, up to the actual policy date.



Table 3: The effect of log density on log price per square meter

|  | (1)                 | (2)                  | (3)                  | (4)                  | (5)                  | (6)                  |
|--|---------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
|  | OLS                 | 2SLS                 | OLS                  | 2SLS                 | OLS                  | 2SLS                 |
| Log density                              | -0.066*<br>(0.035)  | -0.815***<br>(0.273) | -0.070***<br>(0.026) | -0.305**<br>(0.132)  | -0.033<br>(0.027)    | -0.481***<br>(0.175) |
| Constant                                 | 9.597***<br>(0.123) | 7.093***<br>(0.908)  | 11.310***<br>(0.345) | 10.247***<br>(0.705) | 11.246***<br>(0.368) | 9.133***<br>(0.948)  |
| Unit characteristics                     | Yes                 | Yes                  | Yes                  | Yes                  | Yes                  | Yes                  |
| Development characteristics              | No                  | No                   | Yes                  | Yes                  | Yes                  | Yes                  |
| Postal district fixed effects            | Yes                 | Yes                  | Yes                  | Yes                  | Yes                  | Yes                  |
| Sales year fixed effects                 | Yes                 | Yes                  | Yes                  | Yes                  | Yes                  | Yes                  |
| Postal district $\times$ sales year f.e. | No                  | No                   | No                   | No                   | Yes                  | Yes                  |
| Weak IV F-statistic                      |                     | 16.98                |                      | 30.2                 |                      | 21.37                |
| R-squared                                | 0.831               |                      | 0.862                |                      | 0.880                |                      |
| Observations                             | 20,713              | 20,713               | 20,713               | 20,713               | 20,713               | 20,713               |

Notes: \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ . Density and price per square meter are in log. Unit characteristics include floor fixed effects and a quadratic in unit area in square meters. Development characteristics include a 99-year leasehold dummy, a special-neighborhood dummy, log distance to the CBD, log distance to the nearest MRT station, the number of facilities in the development (e.g., gym, pool, etc.), a dummy variable for whether the development was CONQUAS-assessed, and an interaction between the CONQUAS-assessed dummy and CONQUAS score. All regressions include sales year fixed effects and postal district fixed effects. Columns 5 and 6 also include interactions between sales year and postal district fixed effects. Standard errors, in parentheses, are clustered by postal code.

Table 4: The effect of log density on log price per square meter, geographically restricted sample

|  | (1)                 | (2)                  | (3)                  | (4)                  | (5)                  | (6)                  |
|--|---------------------|----------------------|----------------------|----------------------|----------------------|----------------------|
|  | OLS                 | 2SLS                 | OLS                  | 2SLS                 | OLS                  | 2SLS                 |
| Log density                              | -0.008<br>(0.030)   | -0.739***<br>(0.205) | -0.002<br>(0.029)    | -0.370***<br>(0.120) | 0.003<br>(0.032)     | -0.506***<br>(0.175) |
| Constant                                 | 9.783***<br>(0.109) | 7.357***<br>(0.676)  | 11.599***<br>(0.379) | 10.286***<br>(0.567) | 11.388***<br>(0.407) | 9.507***<br>(0.757)  |
| Unit characteristics                     | Yes                 | Yes                  | Yes                  | Yes                  | Yes                  | Yes                  |
| Development characteristics              | No                  | No                   | Yes                  | Yes                  | Yes                  | Yes                  |
| Postal district fixed effects            | Yes                 | Yes                  | Yes                  | Yes                  | Yes                  | Yes                  |
| Sales year fixed effects                 | Yes                 | Yes                  | Yes                  | Yes                  | Yes                  | Yes                  |
| Postal district $\times$ sales year f.e. | No                  | No                   | No                   | No                   | Yes                  | Yes                  |
| Weak IV F-statistic                      |                     | 28.17                |                      | 37.66                |                      | 21.94                |
| R-squared                                | 0.811               | 0.569                | 0.847                | 0.802                | 0.859                | 0.795                |
| Observations                             | 16,215              | 16,215               | 16,215               | 16,215               | 16,215               | 16,215               |

Notes: \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ . This restricted sample includes only transactions from developments in postal districts that contain at least one 1.4 FAR development and one 1.6 to 2.1 FAR development. Density and price per square meter are in log. Unit characteristics include floor fixed effects and a quadratic in unit area in square meters. Development characteristics include a 99-year leasehold dummy, a special-neighborhood dummy, log distance to the CBD, log distance to the nearest MRT station, the number of facilities in the development (e.g., gym, pool, etc.), a dummy variable for whether the development was CONQUAS-assessed, and an interaction between the CONQUAS-assessed dummy and CONQUAS score. All regressions include sales year fixed effects and postal district fixed effects. Columns 5 and 6 also include interactions between sales year and postal district fixed effects. Standard errors, in parentheses, are clustered by postal code.

Table 5: Shoebox and non-shoebox sub-samples

|                               | (1)<br>Shoebox       | (2)<br>Non-shoebox   |
|-------------------------------|----------------------|----------------------|
| Log density                   | -0.818**<br>(0.386)  | -0.304**<br>(0.151)  |
| Constant                      | 12.173***<br>(0.800) | 10.419***<br>(0.770) |
| Unit characteristics          | Yes                  | Yes                  |
| Development characteristics   | Yes                  | Yes                  |
| Postal district fixed effects | Yes                  | Yes                  |
| Sales year fixed effects      | Yes                  | Yes                  |
| Weak IV F-statistic           | 5.32                 | 28.61                |
| Observations                  | 4,066                | 16,647               |

Notes: \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ . Density and price per square meter are in log. Shoebox units are units with 50 square meters or less of floor area. Unit characteristics include floor fixed effects and a quadratic in unit area in square meters. Development characteristics include a 99-year leasehold dummy, a special-neighborhood dummy, log distance to the CBD, log distance to the nearest MRT station, the number of facilities in the development (e.g., gym, pool, etc.), a dummy variable for whether the development was CONQUAS-assessed, and an interaction between the CONQUAS-assessed dummy and CONQUAS score. All regressions include sales year fixed effects and postal district fixed effects. Standard errors, in parentheses, are clustered by postal code.

Table 6: The effect of development composition on log price per square meter

|                                  | (1)                  | (2)                  | (3)                  |
|----------------------------------|----------------------|----------------------|----------------------|
|                                  | 2SLS                 | 2SLS                 | 2SLS                 |
| Percentage of shoebox units      | -0.003**<br>(0.002)  |                      |                      |
| Percentage of medium-sized units |                      | -0.006<br>(0.004)    |                      |
| Percentage of large units        |                      |                      | 0.002**<br>(0.001)   |
| Constant                         | 11.410***<br>(0.424) | 11.704***<br>(0.445) | 11.442***<br>(0.359) |
| Unit characteristics             | Yes                  | Yes                  | Yes                  |
| Development characteristics      | Yes                  | Yes                  | Yes                  |
| Postal district fixed effects    | Yes                  | Yes                  | Yes                  |
| Sales year fixed effects         | Yes                  | Yes                  | Yes                  |
| Weak IV F-statistic              | 20.17                | 2.76                 | 57.27                |
| Observations                     | 20,713               | 20,713               | 20,713               |

Notes: \*\*\*  $p < 0.01$ , \*\*  $p < 0.05$ , \*  $p < 0.1$ . Density and price per square meter are in log. Shoebox units are units with 50 square meters or less of floor area, medium-sized units have more than 50 square meters up to 106 square meters (the 75<sup>th</sup> percentile), and large units have more than 106 square meters. Unit characteristics include floor fixed effects and a quadratic in unit area in square meters. Development characteristics include a 99-year leasehold dummy, a special-neighborhood dummy, log distance to the CBD, log distance to the nearest MRT station, the number of facilities in the development (e.g., gym, pool, etc.), a dummy variable for whether the development was CONQUAS-assessed, and an interaction between the CONQUAS-assessed dummy and CONQUAS score. All regressions include sales year fixed effects and postal district fixed effects. Standard errors, in parentheses, are clustered by postal code.

Table 7: Average price per  $m^2$  before and after the policy change for the average 1.4 FAR development

|                           | Area ( $m^2$ ) | Proportion of total units |       | Log density |        | Log price per $m^2$ |       | Price per $m^2$ |          |
|---------------------------|----------------|---------------------------|-------|-------------|--------|---------------------|-------|-----------------|----------|
|                           | (1)            | (2)                       | (3)   | (4)         | (5)    | (6)                 | (7)   | (8)             | (9)      |
|                           |                | Before                    | After | Before      | After  | Before              | After | Before          | After    |
| <u>A. Table 3, col. 4</u> |                |                           |       |             |        |                     |       |                 |          |
| Shoebox                   | 42.24          | 0.313                     | 0.170 | -3.958      | -4.081 | 9.037               | 9.075 | 8,454.97        | 8,778.16 |
| Medium-sized              | 77.12          | 0.486                     | 0.572 | -3.958      | -4.081 | 8.923               | 8.961 | 7,544.59        | 7,832.99 |
| Large                     | 145.28         | 0.202                     | 0.258 | -3.958      | -4.081 | 8.750               | 8.787 | 6,342.47        | 6,584.92 |
| Ave. price                |                |                           |       |             |        |                     |       | 7,587.13        | 7,671.47 |
| <u>B. Table 3, col. 6</u> |                |                           |       |             |        |                     |       |                 |          |
| Shoebox                   | 42.24          | 0.313                     | 0.170 | -3.958      | -4.081 | 8.995               | 9.054 | 8,105.35        | 8,599.57 |
| Medium-sized              | 77.12          | 0.486                     | 0.572 | -3.958      | -4.081 | 8.873               | 8.933 | 7,180.25        | 7,618.06 |
| Large                     | 145.28         | 0.202                     | 0.258 | -3.958      | -4.081 | 8.687               | 8.746 | 5,959.81        | 6,323.21 |
| Ave. price                |                |                           |       |             |        |                     |       | 7,223.70        | 7,450.64 |

Shoebox units are units with 50 square meters or less of floor area, medium-sized units have more than 50 square meters up to 106 square meters (the 75<sup>th</sup> percentile), and large units have more than 106 square meters. Column (1) contains the mean area in square meters of shoebox, medium-sized, and large units of 1.4 FAR developments. Columns (2) and (3) contain the development proportion of the three unit types, before and after the regulation. Columns (4) and (5) contain the mean log density of 1.4 FAR developments, before and after the regulation. Columns (6) and (7) contain the predicted log price per square meter using the 2SLS regression estimates in Table 3, column (4) (panel A) and in Table 3, column (6) (panel B) for units with column (1) area, in postal district 17, on the 3<sup>rd</sup> floor, and sold in 2012, and log density reported in columns (4) and (5), respectively. In columns (8) and (9), we transform the predicted log price in columns (6) and (7) to price per square meter using the formula  $\hat{p} = \exp(\hat{\sigma}^2/2) \exp(\widehat{\ln p})$ , where  $\sigma^2$  is the variance of the error term. Average price in the last row of columns (8) and (9) in each panel is the weighted average price per square meter using price in columns (8) and (9) and the shares in columns (2) and (3).

Table 8: Average price per  $m^2$  before and after the policy change for the average 1.4 FAR development

|            | Size ( $m^2$ ) | Proportion of total units |       | Log density |        | Log price per $m^2$ |       | Price per $m^2$ |          |
|------------|----------------|---------------------------|-------|-------------|--------|---------------------|-------|-----------------|----------|
|            | (1)            | (2)                       | (3)   | (4)         | (5)    | (6)                 | (7)   | (8)             | (9)      |
|            |                | Before                    | After | Before      | After  | Before              | After | Before          | After    |
| Decile 1   | 37.85          | 0.150                     | 0.034 | -3.958      | -4.081 | 9.053               | 9.09  | 8,587.33        | 8,915.59 |
| Decile 2   | 43.43          | 0.078                     | 0.115 | -3.958      | -4.081 | 9.033               | 9.07  | 8,419.89        | 8,741.74 |
| Decile 3   | 49.20          | 0.106                     | 0.034 | -3.958      | -4.081 | 9.013               | 9.051 | 8,253.80        | 8,569.30 |
| Decile 4   | 56.95          | 0.104                     | 0.102 | -3.958      | -4.081 | 8.987               | 9.025 | 8,041.80        | 8,349.20 |
| Decile 5   | 69.56          | 0.087                     | 0.164 | -3.958      | -4.081 | 8.946               | 8.984 | 7,722.06        | 8,017.24 |
| Decile 6   | 77.86          | 0.094                     | 0.090 | -3.958      | -4.081 | 8.921               | 8.959 | 7,527.92        | 7,815.67 |
| Decile 7   | 89.35          | 0.102                     | 0.090 | -3.958      | -4.081 | 8.887               | 8.925 | 7,278.49        | 7,556.72 |
| Decile 8   | 103.24         | 0.098                     | 0.145 | -3.958      | -4.081 | 8.849               | 8.887 | 7,005.27        | 7,273.05 |
| Decile 9   | 121.46         | 0.086                     | 0.108 | -3.958      | -4.081 | 8.803               | 8.84  | 6,689.55        | 6,945.26 |
| Decile 10  | 175.55         | 0.095                     | 0.117 | -3.958      | -4.081 | 8.693               | 8.731 | 5,995.29        | 6,224.46 |
| Ave. price |                |                           |       |             |        |                     |       | 7,609.79        | 7,690.35 |

Deciles of floor area are computed using transactions in 1.4 FAR developments. Column (1) contains the mean area in square meters by unit area decile. Columns (2) and (3) contain the development proportion of the three unit types, before and after the regulation. Columns (4) and (5) contain the mean log density of 1.4 FAR developments, before and after the regulation. Columns (6) and (7) contain the predicted log price per square meter using the 2SLS regression estimates in Table 3, column (4) (panel A) and in Table 3, column (6) (panel B) for units with column (1) area, in postal district 17, on the 3<sup>rd</sup> floor, and sold in 2012, and log density reported in columns (4) and (5), respectively. In columns (8) and (9), we transform the predicted log price in columns (6) and (7) to price per square meter using the formula  $\hat{p} = \exp(\hat{\sigma}^2/2) \exp(\widehat{\ln p})$ , where  $\sigma^2$  is the variance of the error term. Average price in the last row of columns (8) and (9) in each panel is the weighted average price per square meter using price in columns (8) and (9) and the shares in columns (2) and (3).

Figure 1: FAR 1.4 and FAR 1.6 to 2.1 developments with postal district boundaries

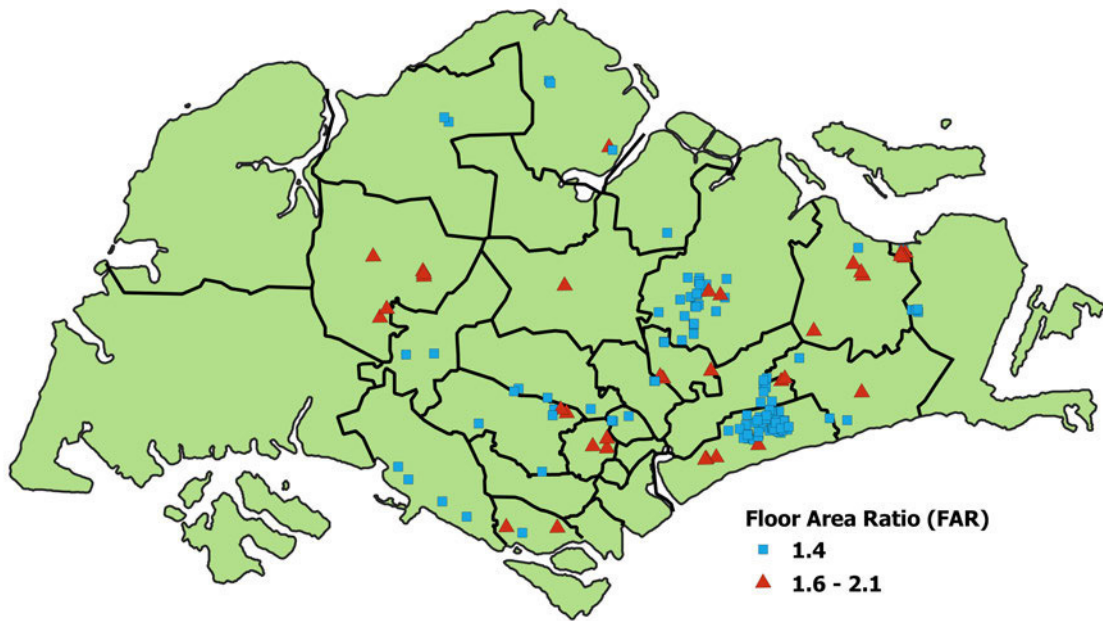


Figure 2: Development density over time by Floor Area Ratio (FAR)

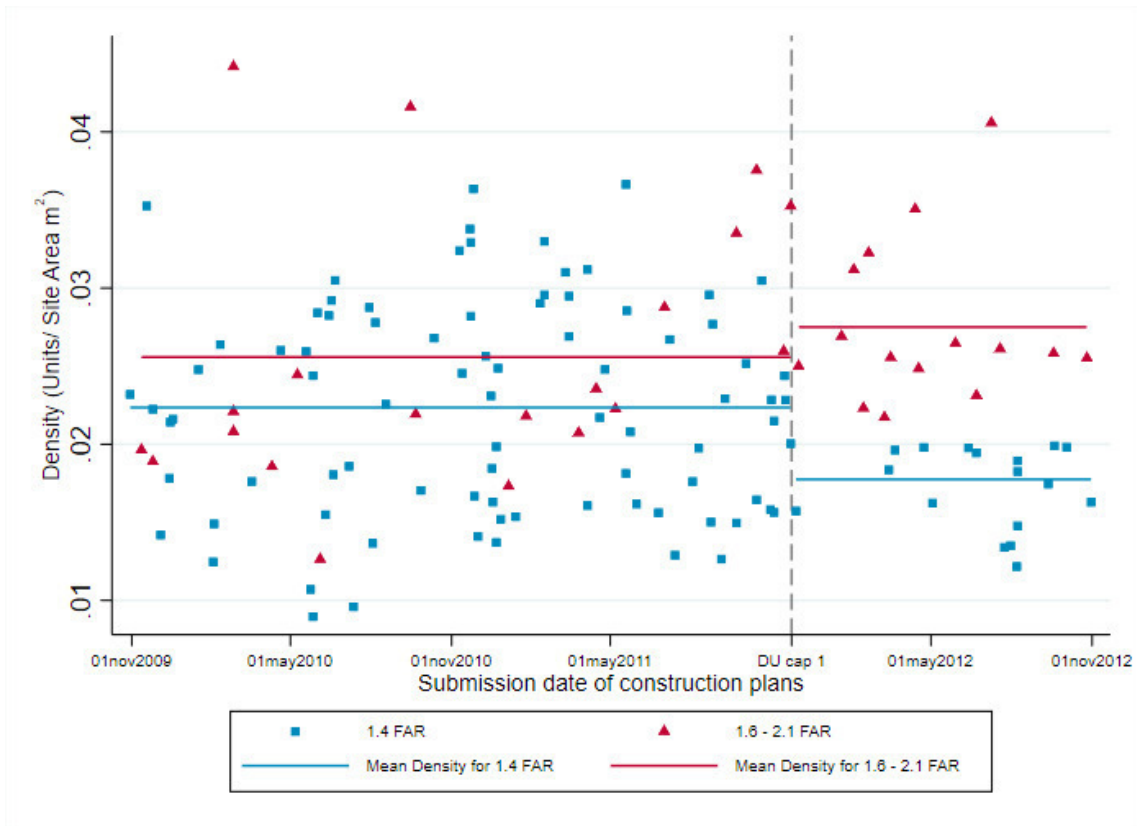
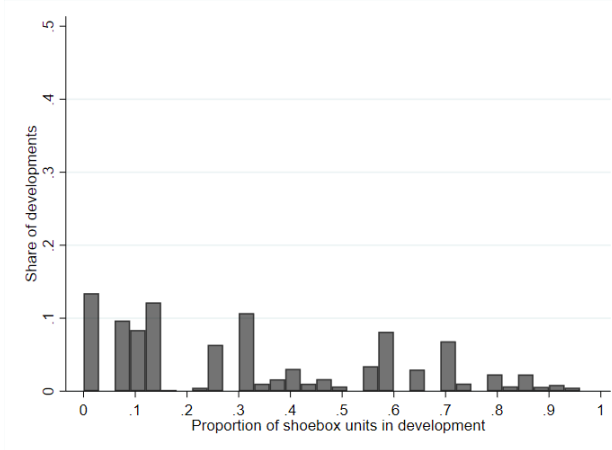


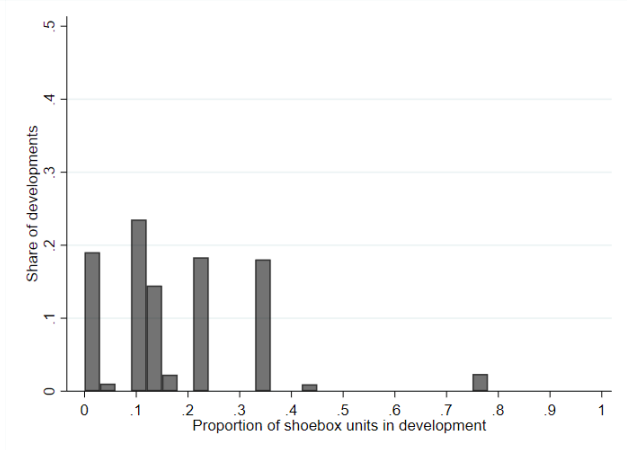


Figure 3: Distribution of shoebox units by FAR, before and after policy change

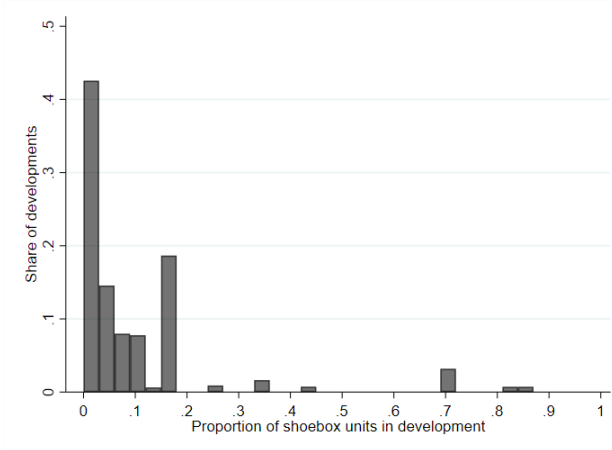
(a) 1.4 FAR developments, before policy change



(b) 1.4 FAR developments, after policy change



(c) 1.6 - 2.1 FAR developments, before policy change



(d) 1.6 - 2.1 FAR developments, after policy change

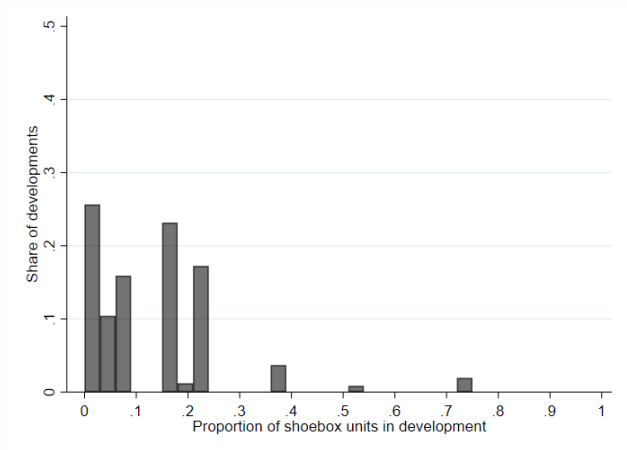
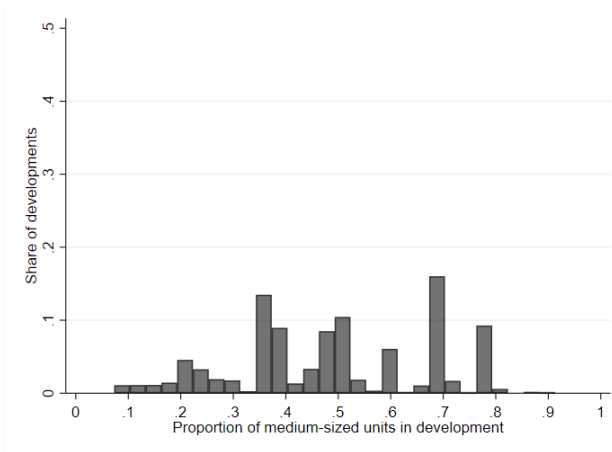
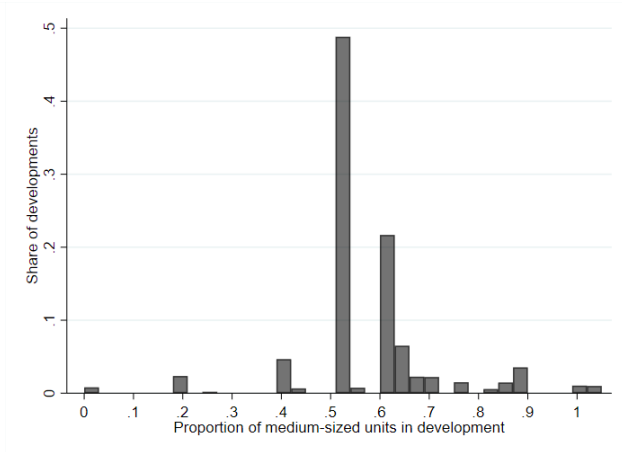


Figure 4: Distribution of medium-sized units by FAR, before and after policy change

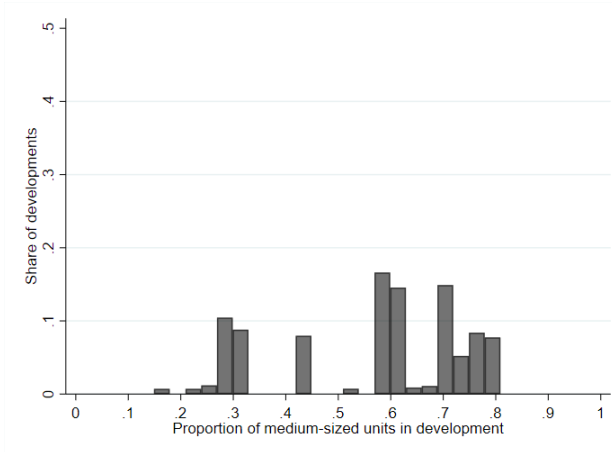
(a) 1.4 FAR developments, before policy change



(b) 1.4 FAR developments, after policy change



(c) 1.6 - 2.1 FAR developments, before policy change



(d) 1.6 - 2.1 FAR developments, after policy change

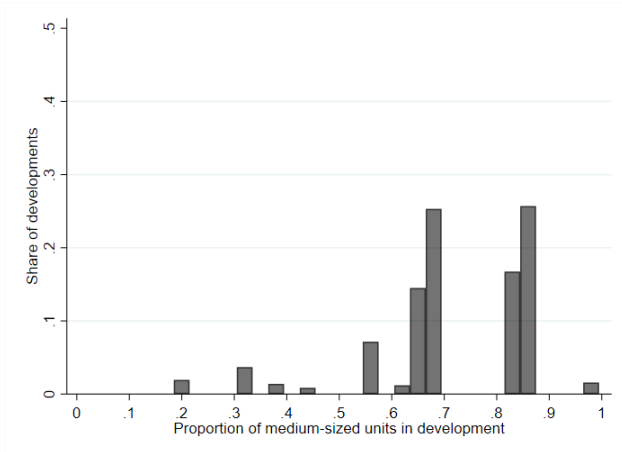
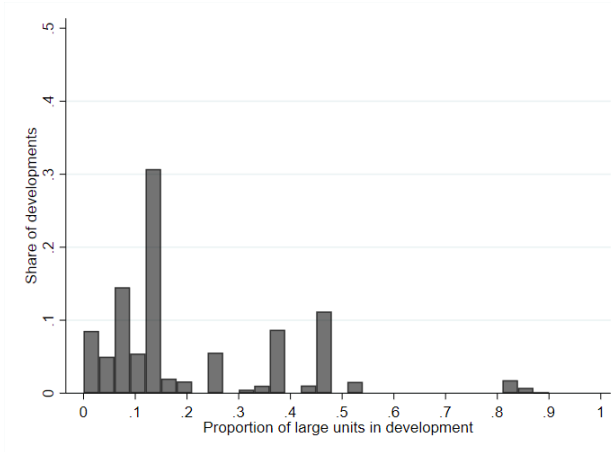
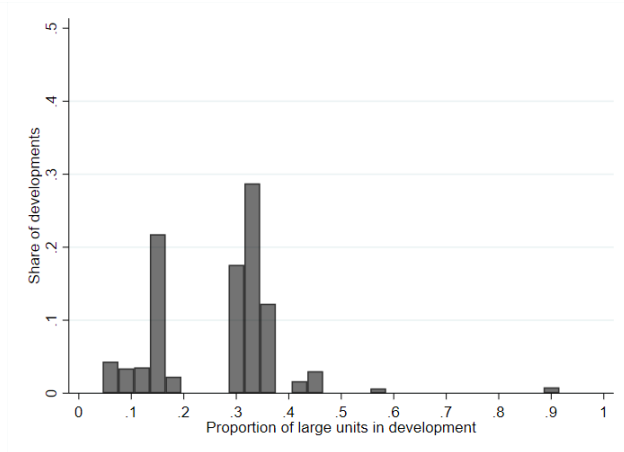


Figure 5: Distribution of large units by FAR, before and after policy change

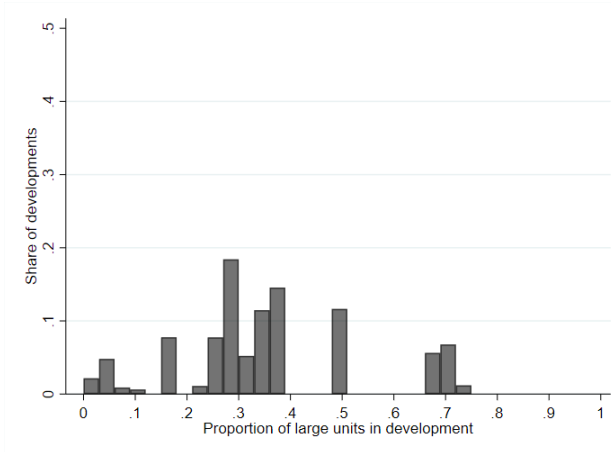
(a) 1.4 FAR developments, before policy change



(b) 1.4 FAR developments, after policy change



(c) 1.6 - 2.1 FAR developments, before policy change



(d) 1.6 - 2.1 FAR developments, after policy change

