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Zuzana Fungáčová, Koen Schoors, Laura Solanko and Laurent Weill

Political cycles and bank lending in Russia



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Zuzana Fungáčová, Koen Schoors, Laura Solanko and Laurent Weill

Political cycles and bank lending in Russia

Abstract

State-owned banks tend to increase lending before elections for the purpose of boosting the reelection odds of incumbent politicians. We employ monthly data on individual banks to study
whether Russian banks increased their lending before presidential elections during 2004–2019, a
period covering four presidential elections. In contrast to the literature, we find that both state-owned
and private banks increased their lending before presidential elections. This result stands for all
loans, as well as separately for firm and household loans. The pre-election lending surge is followed
by a deterioration of loan quality the following year, indicating the lending increase was not driven
by higher growth prospects or some positive economic shock. The effect is substantially greater for
large banks and banks more involved in lending activities. Our main finding that all types of banks
in Russia increase their lending before presidential elections supports the view that the authorities
in an electoral autocracy like Russia can influence lending of both private and state-owned banks
for political reasons.

JEL Codes: G21, P34.

Keywords: bank, lending, politics, Russia.

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1 Introduction

Ever since Vladimir Putin assumed the Russia's presidency in 2000, there have been spirited discussions about political intervention in the electoral process. The literature on Russia investigates a number of channels available for achieving political results, including media control (Enikolopov, Petrova and Zhuravskaya, 2011) and electoral fraud (Klimek et al., 2012). It further suggests that the authorities take an active interest in shaping the outcomes of elections.

Banks potentially constitute a channel for influencing election outcomes in Russia. Recent literature drawing from experience in several countries stresses the relation between bank lending and elections, showing that state-owned banks can be utilized to influence the electoral outcome. Sapienza (2004) explains that the government can utilize state-owned banks to pursue its own interests such as improving their chances for reelection or avoiding political instability. Dinc (2005) finds that state-owned banks increase lending in election years relative to private banks in major emerging markets, a finding that implies state-owned bank lending can be used to influence political outcomes. Carvalho (2014) shows the impact of lending of state-owned banks on the decision-making of Brazilian firms aligns with electoral outcomes. He observes that state-owned bank lending is associated with employment growth by firms in politically attractive regions near elections. Englmaier and Stowasser (2017) find that German savings banks, where local politicians are involved in their management, adjust their lending policies in response to local electoral cycles. For Turkey, Bircan and Saka (2019) show that, unlike private banks, state banks engage in politically motivated lending during the lead-up to local elections.

We investigate whether banks increase their lending before presidential elections in Russia utilizing detailed monthly data on individual banks from the Central Bank of Russia (CBR) to identify with high precision any changes in bank lending before and after elections. Our observation period runs from 2004 to 5/2019, a period that includes four presidential elections (2004, 2008, 2012 and 2018).

In line with the empirical literature, we start by testing whether Russian state-owned banks boosted their lending before presidential elections. In his seminal paper, Dinc (2005) suggests a motivation for this behavior, noting that politicians in some countries go so far as to use their power to influence private bank lending as well. We thus proceed by verifying empirically whether all Russian banks, state-owned and private, increase their lending before presidential elections. This is hardly unexpected behavior in Russia's case as the interplay between politics and banking in an electoral autocracy need not be restricted to state-owned banks.

¹ This conjecture was later raised for Turkey by Baum, Caglayan and Talavera (2010).

Even with a bias in favor of incumbents, electoral autocracies go through the motions of organizing competitive elections in order to bolster the legitimacy of the autocrat. As noted by Egorov and Sonin (2018), the election outcome, combined with the leader's decision to allow opposition candidates to compete, signals the strength of the non-democratic leader and can be utilized to dissuade protest against the regime. The incumbent leader has a number of tools available for intervening in the electoral process to influence the electoral outcome. Besides propaganda and fraud, they include pressuring banks to enhance lending ahead of an election. The application of such duress has been observed for state-owned banks in developed and emerging market economies.

Electoral autocracies differ from democracies and traditional autocracies in the sense that these regimes are characterized by the most suitable combination of incentives and possibilities to influence both state-owned and private banks. On contrast, democracies have fewer opportunities to influence elections by pressuring private banks due to checks and balances and the presence of a free press willing to give the government a hard time if it puts too much pressure on private banks. Similarly, there is little incentive to pressure banks in traditional autocracies as such leaders have little concern about influencing elections. There is no electoral franchise or elections with credible outcomes.

Our key hypothesis states generally that there is political interference with bank lending policies prior to presidential elections in Russia. More specifically, we test the hypothesis that the authorities are able to motivate all banks to boost their lending during the run-up to a presidential election. Evidence of greater lending prior to presidential elections (simply elections hereafter) may also be the result of a higher demand for loans, however, and not a supply-related factor such as political pressure.²

To make sure our results are not driven by demand-related considerations such as a preelection economic expansion, we control for macroeconomic fluctuations in our estimations. We also determine whether loan quality deteriorated in the first post-election year relative to other years. If a pre-election lending is driven solely by economic factors, there is no reason for a significant rise in bad loans after elections. The deterioration of loan quality after elections, however, accords with the view that the pre-election lending surge is predominantly unrelated to economic factors. This finding would bolster our interpretation of political interference.

There are several reasons why banks might want to boost lending ahead of elections. First, the authorities could encourage them to do so. If the authorities directly or indirectly animate banks to increase lending, we assume such political inference focuses on banks with the largest potential

² It is unsurprising that we find no documentation confirming such pressure. Its absence is fully in line with the thesis that electoral autocracies seek to influence elections covertly.

impact. We thus check to see whether large banks and banks with a high share of loans on their balance sheet are more likely to boost lending prior to an election. Such a finding would accord with the view that political influence targets banks where a lending surge has the most expected impact on the election outcome. If the increase in lending was driven by some other phenomenon than the pressure from the authorities, we would expect to see a similar increase in lending by all banks as an election approaches, not just banks most involved in lending.

Banks may also have their own incentives for boosting lending before an election. For example, banks with weak fundamentals could voluntarily decide to increase their supply of loans in order to buy implicit protection or even regulatory forbearance in the face of the risk of imminent license withdrawal. Thus, we also test to see if the pre-election lending of banks with low capital ratios and poor loan quality differs from the lending behavior of other banks.

Additionally, the supply of loans may be increased in a different way for different types of loans. Our data allow us to distinguish between loans to firms and loans to households and check if any of those increases more before elections. While all lending enhances economic activity over the short run and pleases borrowers, loans to firms can be used in Russia to get employers to exert pressure in the workplace, which has been shown to be a key site of political mobilization in Russia (Frye, Reuter and Szakonyi, 2014).

Overall, the following analysis contributes to our understanding the electoral successes of Putin in presidential elections by scrutinizing the influence of bank lending in this political outcome. It also adds to the literature on banks influencing political outcomes. While former papers have considered democracies (for Brazil: Carvalho, 2014; for Germany: Englmaier and Stowasser, 2017) or cross-country samples (Dinc, 2005), we extend this literature by considering how bank lending can be used to affect electoral outcome in an electoral autocracy. We also extend this literature by considering the post-election effects of this influence as we analyze the effects of bank lending in electoral times on bad loans in the following years.

The rest of the article is structured as follows. Section 2 discusses the development of Russian banking sector during the last decades. Section 3 presents the data and methodology. Section 4 displays the results. Section 5 concludes.

2 Presidential elections and the Russian banking sector

After the turbulent 1990s, the regime and the new president elected in March 2000 succeeded to balance the budget, to re-monetize the economy and to pursue liberally-minded structural reforms. Macroeconomic stabilization supported economic growth and prepared the ground for rapid development of the banking sector in the second half of the decade.

A new law on presidential elections was approved in 2003,³ and presidential elections were held as scheduled in March 2004. Increasing oil prices supported public finances, and both portfolio investments and FDI flooded into the Russian economy. As the Central Bank of Russia kept the nominal exchange rate stable, real appreciation took place through higher domestic inflation. A deposit insurance scheme was put in place in 2004, and the last remaining restrictions on capital account were removed in 2006. The overall economic boom fostered growth in modern banking. In early 2006, Russia had 1,244 operating banks, most of them tiny, almost always unlisted and owned by a handful of wealthy individuals. Banking sector assets to GDP increased from just 40 % in 2004 to 60 % in late 2007, when the global financial crisis hit Russia. Despite a looming depression, presidential elections were again held as scheduled in March 2008.

Generous state support and temporary relaxation of regulatory measures helped Russia's banking sector weather the global financial crisis relatively unscathed. The sector remained fragmented, however. As a legacy of the 1990s, Russia had over a thousand banks, but was essentially dominated by a few state-controlled universal banks. While the crisis further increased the share of state-owned banks enjoying de facto state guarantees, it also made all banks more reliant on the state as a source of funds. Many banking sector support measures included increased state guidance on lending priorities. The role of the state increased in other sectors as well. Centralization of public finances increased the powers of federal authorities. The regime started to foster consolidation of state-owned companies in larger conglomerates and state corporations enjoying special administrative privileges.

The role of foreign banks remained minor. Operating foreign banks could be broadly classified into three categories: banks registered abroad but mainly controlled by Russian individuals, foreign banks competing with Russian banks in the domestic market and foreign banks serving foreign-owned corporate clients. The last group largely vanished in the 2010s. The group of genuinely foreign banks included perhaps a half-dozen financial institutions. Their combined market share remained at around 10 % of total lending.

The central bank consolidated under its roof all financial markets supervisory functions, and launched a determined process of cleaning up the banking sector in 2013. The new supervisory body enjoyed a clear mandate to weed out the weakest and most obscure financial institutions.

³ The president is elected for a four-year period and the elections are to be held in the same month as the previous elections were held. The law was amended in 2008 to increase the term to six years. For a thorough description of elections laws and practices, see e.g. OSCE election observation monitoring reports available at https://www.osce.org/odihr/elections/russia.

⁴ This structure does not enable us to use foreign banks as a separate group in our analysis.

A new crisis hit the banking sector in 2014–2015 in the form of an oil price collapse and Western sanctions on Russia. Faced with a stubborn economic recession, the monetary policy framework was overhauled dramatically in late 2014 as the central bank shifted to inflation targeting. The ruble was allowed to float freely, leading to a sizable depreciation (and as a consequence, an increase in ruble-denominated value of bank's loan books as about 20 % of lending was in foreign currencies).

The 2014–2015 crisis intensified the clean-up of the banking sector. The number of operating credit institutions dropped from 955 at the end of 2012 to 619 at the end of 2016 and 442 at the end of 2019. Additionally, a number of faltering top-50 banks were taken over by the CBR in the latter half of 2017, and many more were assigned to the Deposit Insurance Authority for rehabilitation. Despite the decreasing number of credit institutions, the trend in bank lending was growing over the whole period under investigation (Figure 1).



Figure 1 Banking sector assets and lending

Source: Central Bank of Russia.

Note: As bank assets and lending are reported in nominal rubles, the sudden peaks in the figure are caused by sudden devaluations increasing the ruble value of foreign-currency denominated assets.

3 Data and methodology

We employ monthly bank-level financial statement data for Russian banks from the Central Bank of Russia (CBR). The period covered starts in 2004 and ends in mid-2019. The dataset contains detailed information on the various kinds of loans provided by banks. All Russian banks are included in the dataset, so there is no selection bias. We combine the data with the information on the bank

ownership provided by Karas and Vernikov (2016), the Central Bank of Russia and individual bank websites. The data concerning the output index for key economic activities come from Rosstat.

To avoid the extreme values, we winsorize the variables included in the analysis at 1 % and 99 % levels. The final sample we use for our estimations constitutes an unbalanced panel of over 131,000 bank-month observations for 1,209 banks. Table 1 describes the summary statistics.

We investigate the change in lending around elections by estimating the following panel regression:

 $loangrowth_{i,t}$

$$= \alpha + \beta * Elections_t + \eta * State - owned_{i,t} + \rho * Elections_t$$

$$* State - owned_{i,t} + \gamma * X_{i,t-1} + \eta * Output \ index_{t-1} + \omega_i + \tau_t + \varepsilon_{i,t}$$

The dependent variable is month-on-month growth in bank lending ($loangrowth_{i,t}$). The main variable of interest is the dummy variable concerning the election period (Elections). It is defined in several alternative ways. First, it is equal to one in months when presidential elections took place, i.e. March 2004, March 2008, March 2012 and March 2018. In the alternative specifications, it is equal to one for the months before these elections. We consider the preceding months separately (February, January, December) or as a pre-election period of three or six months. These alternative definitions of the pre-election period yield eight different specifications of our baseline model.

We also include a dummy variable for state ownership (*State-owned*) and an interaction variable between *Elections* and *State-owned*. The coefficient on the interaction term *Elections*×*State-owned* directly tests the hypothesis that the pre-election lending behavior of state-owned banks differs from the pre-election behavior of private banks.

We control for several bank-specific variables that have been shown to influence loan growth. These include bank size defined by logarithm of total assets, capital ratio, the ratio of bad loans to loans, as well as a loans-to-assets ratio to account for the business model of the bank. All these control variables are lagged by one period. We also account for macroeconomic development by including the output index for key economic activities. All specifications are estimated using a random effects model with robust standard errors. We include month fixed effects.

4 Results

This section presents our results for the link between elections and bank lending. We start with our main estimations to investigate whether prior to elections the lending behavior of state-owned banks

differs from private banks. We then test whether changes in bank lending before elections are supply- or demand-driven. The last subsection discusses our robustness checks.

4.1 Main estimations

The main estimations explaining changes in total lending are reported in Table 2. For robustness, we consider several definitions of the election period. Since elections take place at different times in March, we consider each of the four months before elections (from December to March). We also test two three-month periods (December–February and January–March) and two six-month periods (September–February and October–March) as pre-election periods.

Our key finding is a positive and significant coefficient for the *Elections* variable in most specifications. With the exception of January⁵, we always find that lending increases before elections regardless of the election period tested. Our conclusion that lending of Russian banks on average increased before elections is in line with the hypothesis that the authorities influence banks, irrespective of ownership, to increase their lending during these periods.

We next investigate whether the increase in lending before elections is more pronounced for state-owned banks than private banks. On one hand, we expect that lending should be enhanced for state-owned banks in line with the literature that shows state-owned banks are exploited to influence political outcomes. On the other hand, state-owned banks and private banks in electoral autocracies like Russia can similarly be influenced.

We observe that the interaction term *Elections*×*State-owned* is not significant for five specifications (September–February, October–March, December, February, March), and significantly positive for three specifications (December–February, January–March, January). Since the hypothesis that state-owned banks increase their lending more than other banks is rejected in five out of eight specifications, we may conclude there is no (or at best very weak and inconsistent) evidence that state-owned banks would increase lending more than private banks before elections.

The estimated coefficients of the control variables are all significant and have the expected sign. *Capital/assets* is positively related to loan growth, which is consistent with the view that higher capitalization supports lending. The coefficient for *Bad loans/loans* is negative, as lower quality of the loan portfolio reduces loan growth. *Size* and *Loans/assets* both reduce lending growth as observed by Bertay, Demirgüç-Kunt and Huizinga (2015). This can be explained by the fact that the ability to increase loans in relative terms is lower for large banks or banks with a high share of loans on their balance sheets. Finally, the coefficient for the variable *Output index* is positive, which accords with the view that economic growth is related to loan growth.

⁵ A possible reason might be the long holidays in the first half of January.

Our main estimations confirm a general bank lending surge prior to presidential elections in Russia. Our hypothesis, which draws on the literature, is that the authorities influence banks to increase lending before elections. We proceed by testing whether the pre-election lending surge could have been driven by other supply or demand factors in order to rule out alternative hypotheses that could also comport with our empirical results.

4.2 Is the increase in lending driven by loan demand?

A surge in lending before elections could also be the consequence of exogenous events influencing the whole economy. For instance, the government could push for specific economic policies that fuel supply, demand or both prior to elections (possibly for political reasons). Such measures would increase credit demand and thereby lending. In such case, our interpretation that the government influences lending of all banks before elections would be incorrect since higher bank lending was not the direct consequence of government behavior but rather the market response to higher economic activity. To account for this, we include the output index in all the estimated specifications. As we focus only on one country, all macroeconomic variables are the same for all banks and the effects of these variables are accounted for by the output index and month dummy variables.⁶

To test the relevance of our interpretation, we perform an additional set of estimations in which we investigate whether deteriorating loan quality follows elections. Our reasoning behind this test is straightforward and in line with Englemaier and Stowasser (2017). If a pre-election surge in lending is driven by economic factors such as increased demand, there is no reason for a sharp rise in bad loans after the elections. Banks would grant more loans before elections in line with prudent financial criteria designed to prevent deterioration of the average quality of their loan portfolio. The observation of substantial increases in bad loans after elections, however, aligns with the view that the pre-election lending surge was caused by non-economic factors such as political interference. That would explain why these loans, for which the decision to grant was not based on financial criteria, are riskier and therefore lead to a deterioration of loan quality and implied higher growth of bad loans.

Specifically, we substitute the dependent variable of loan growth with the yearly growth of bad loans in our main equation. Since we aim to analyze whether bad loans increase after presidential elections and recognize that bad loans do not come to light immediately, we consider a one-year post-election period. This should be long enough for a deterioration in loan quality to manifest, but not too long to decouple causation from the presidential election. In the following set of results,

⁶ Furthermore, additional estimations include (one at the time): monthly federal budget expenditures, monthly average corporate lending rate (RUB denominated, up to 1 year), monthly EPU uncertainty index and the interbank rate (1-day MIACR monthly average). Our main results stay unchanged. Full results are available upon request.

Elections is a dummy variable equal to one if the period corresponds to the year following presidential elections. We consider two definitions of bad loans to test the sensitivity of our results: bad loans and overdue loans (bad loans are recognized as overdue before being classified as bad). In line with the main estimations, we also check the potential influence of state ownership on the results.

Table 3 displays the results of the estimations. We find that *Elections* is significantly positive in all specifications, indicating that bad loans grow faster in post-election years. The interaction term *Elections*×*State-owned* is not significant in the case of bad loans and only marginally significant for overdue loans, supporting the view that the post-election increase of bad loans is not significantly different for state-owned banks in comparison to other banks. These results provide support for our interpretation that the lending surge prior to elections is not strictly driven by economic factors. This approach rules out alternative explanations for our empirical results and corroborates the view that the pre-election surge in loans relates to political influence. The result that bad loan growth increases in the year following presidential elections also suggests the existence of welfare costs associated with this form of political interference.

4.3 Is the pre-election increase in bank lending focused on a specific type of lending or bank?

If the lending surge is motivated by political considerations, it remains an open question whether it is politically optimal to do this across the board for all types of loans, or to focus on certain types of loans where the effect on elections outcomes can be expected to be larger. One underlying motivation for the authorities to increase lending before elections could be to enhance economic activity in the short-run, in this way raising employment, wages and satisfaction and therefore boosting the odds of reelection. Both loans to households (through increased consumption) and loans to corporates would yield this type of effect.

Loans to households have the additional effect of immediate consumer satisfaction because of the marginal utility of increased consumption and may be very useful to sway the opinions of voters. But loans to corporates may also have effects beyond the direct boost of economic activity in the context of Russia. Frye, Reuter and Szakonyi (2014) explain that Russian employers, in particular managers of private firms, influence the voting behavior of their employees and that this type of workplace voter mobilization is common in Russia. A surge in corporate lending prior to elections may be instrumental in incentivizing managers to use this mechanism of voter workplace mobilization. Schoors and Weill (2020), on the other hand, show how specifically Sberbank corporate loans were used to this purpose in the presidential election of March 2000. The workplace voter mobilization effect of additional corporate loans may therefore be limited to lending activity by only one bank and not show up in our framework covering all banks.

There are therefore good reasons to believe that the lending surge would occur in both corporate and household lending and that it would not be necessarily more pronounced in corporate lending than in household lending. To analyze whether our pre-election political influence effect is more pronounced for firm or household loans, we run the main estimations separately for firm loan growth and household loan growth. Tables 4 and 5 report the estimations for firm loans and household loans, respectively.

As hypothesized, we find that lending increases before elections for both firm and household loans. The results are not stronger for corporate loans than household loans. *Elections* is significantly positive in all estimations with the exception of specifications for December and January for firm loans and January for household loans. We additionally observe no difference between state-owned banks and other banks as the interaction *Elections*×*State-owned* is not significant. These results support the view that political interference in pre-election bank lending occurs indiscriminately for firm loans and household loans.

Second, we check to see if lending increased prior to elections for certain types of banks in order to maximize the impact on political outcomes. We expect that political influence efforts to boost lending target those banks most important for lending. Indeed, it is optimal for the authorities to exert more influence on these banks that are the lynchpin of lending activities in Russia rather than to put an equal amount of effort in influencing small banks that only have very small share of loans on their balance sheets.

We therefore analyze whether large banks and banks with a higher share of loans on their balance sheets exhibit larger lending surges in pre-election periods. If the increase in lending is caused by a factor other than political interference, we should observe the same increase in loans for all banks during election times, not just banks that are most important for aggregate lending activity. If our hypothesis of political influence is right, however, the result of higher loan growth before elections should be more pronounced for the banks most involved in lending. To test our hypothesis, we re-estimate our regressions by considering separately large versus small banks and high-lending versus low-lending banks. We classify banks as large or small, based on the size of their balance sheet relative to the median balance sheet of the sample. We define banks as "high-lending" versus "low-lending" based on their loans-to-assets ratio relative to the median loans-to-assets ratio of the sample.

Table 6 reports the estimates for large versus small banks. The coefficient for *Elections* is significantly positive for large banks in all specifications (with just one exception). For small banks, in contrast, the coefficient for *Elections* is only significantly positive for four of the eight specifications (September–February, October–March, December–February and December). The size of the

effect tends to be larger for large banks, suggesting that the effects are generally more pronounced for large banks as hypothesized.

Table 7 displays the estimations for high-lending and low-lending banks. We observe higher loan growth before elections for high-lending banks. The results are consistent with our observations in the above comparison of large and small banks in the sense that results are the same for large banks and high lending banks while they are similar for small banks and low lending banks. The coefficient of *Elections* is again significantly positive for all specifications with only one exception (January) for high lending banks while for low lending banks *Elections* is only significantly positive for five out of eight specifications (September–February, October–March, December–February, December, and March). We thus conclude that lending has particularly increased before elections for the banks most involved in lending, which lends further support to our hypothesis of political interference in bank lending before elections.

Third, banks can differ in their connections to the government, irrespective of ownership type. Therefore, we investigate if loan growth prior to elections is greater for banks with higher shares of government deposits. The underlying assumption is that banks with high shares of government deposits can be incentivized to provide lending prior to elections in order to secure lucrative government business in the future. This is in line with former works showing that politicians reciprocate by granting easier access to government contracts (for the US: Goldman, Rocholl and So, 2013; for South Korea: Schoenherr, 2019; for France: Delatte, Matray and Pinardon-Touati, 2020).

We test this hypothesis by including the dummy variable *Gov deposits*. It is equal to one if the ratio of government deposits to total deposits in August the year before the election (so that we consider the ratio valid for a date before any definition of election period considered in the estimations) is higher than the median, and zero otherwise. We add this variable and its interaction with *Elections* in the estimations so that the interaction term reveals whether banks with a greater share of government deposits increase their lending more prior to elections.

Table 8 reports the results. We observe that the interaction term *Elections*×*Gov deposits* is not significant in any specification. Thus, we can reject the hypothesis that banks with a greater share of government deposits have higher loan growth prior to elections. Access to lucrative government deposits does not seem to influence the lending behavior of banks prior to elections.

Finally, we study whether banks with weak bank fundamentals are more likely to increase their supply of loans before elections. Managers of these banks may aim to increase lending on the expectation of improving their chances of retain their banking license. They essentially buy implicit protection and raise the bank's license value. A large number of banks in Russia have seen their licenses pulled over the past two decades. While decisions by the bank supervision department at

the Central Bank of Russia to withdraw an individual banking license are typically driven by prudential concerns (Claeys and Schoors, 2007), it cannot be ruled out that these decisions are also influenced by political motives.

To investigate this hypothesis, we consider two definitions of weak bank fundamentals based on two criteria: capital ratio (*Capital/assets*) and loan quality (*Bad loans/loans*). Banks with low capital ratios and poor loan quality are most vulnerable to bank license withdrawal and can be assumed to be more prone to behave in a way that is politically desirable. We include alternatively the interaction term of *Elections* with *Capital/assets* and *Bad loans/loans* to check whether banks with weak fundamentals exhibit a different lending behavior prior to elections. If the hypothesis that weak fundamentals influence the behavior of banks prior to elections is valid, we should observe that the interaction term *Elections*Capital/assets* is negative and that the interaction term *Elections*Bad loans/loans* is positive, indicating that banks with poorer capitalization and loan quality increase their lending more prior to elections.

Tables 9 and 10 display the results. We find that the interaction term is not significant in any specification, with one exception for loan quality. We thus find no support for the view that banks with weak fundamentals have higher loan growth prior to elections to improve their chances of holding on to their banking license.

To summarize, the lending surge observed prior to presidential elections cannot be explained by alternative economic factors. We see an across-the-board pattern for corporate loans, household loans, weak banks, strong banks and banks with and without large amounts of government deposits. The only exception to the generality of our effect seems to be that the effect is slightly more pronounced for banks heavily involved in aggregated lending.

Our findings support the hypothesis of general political influence on bank lending prior to elections. The notion that all types of banks, not just state-owned banks, increase their lending in election periods certainly sounds odd for a democratic regime with arms-length relations between business and politics, but in a heavily managed electoral autocracy like Russia's this result should not come as a surprise. Where weak rule of law prevails, the party in power may possess almost unrestricted means to influence any business by a multitude of unobserved administrative means (e.g. changing regulations or modifying their implementation). Private business may well see it in their self-interest to promote good causes brought forward by the party of power and in equilibrium self-select into what is perceived as politically desirable lending behavior.

4.4 Robustness checks

We perform three robustness checks to check the sensitivity of our analysis.

First, our analysis is based on the four presidential elections which have taken place in Russia between 2004 and 2018. To investigate whether our results are driven by any of these four electoral episodes, we redo our main estimations by dropping one election period at a time. This way, we perform four estimations in which there are only three election periods and we can investigate if the results stand.

The results of these estimations are reported in Table 11. We observe the results for excluding each election year at a time are similar to the results obtained for all four presidential elections. The *Elections* variable is positive and significant in most estimations. In the main estimations including all election years, we had a significantly positive coefficient for *Elections* in all eight specifications with the exception of January. Right now, the *Elections* variable is significantly positive in six specifications when 2004 elections are excluded, in seven specifications when 2008 and 2012 elections are excluded, and in all eight specifications when 2018 elections are dropped. Therefore, we conclude that our main finding of a pre-election lending surge holds even when we exclude any of the four election episodes, indicating our findings are very robust to the choice of the election period.

Second, we conduct a placebo test. To this end, we falsify the year of the elections, while preserving the monthly pattern of the *Elections* variable. We assume that the elections have taken place in April 2006, 2010, and 2014, rather than 2004, 2008 and 2012, and redefine the *Elections* variable accordingly. If our results were driven by other unobserved events related to the time period in the year, rather than by the election itself, we should still observe that coefficient for *Election* is significantly positive.

Table 12 reports the estimations. We find that *Elections* is not significantly positive in any estimation. Thus, the placebo test confirms our main finding that banks increase lending in the preelection period only in election years.

Third, as about a quarter of corporate lending on average is denominated in foreign currencies in Russia, currency revaluations can greatly affect the value of the ruble-denominated total loan stock. To check the robustness of our results, we consider only ruble-denominated loans as the main dependent variable. Our main results stay unchanged. These results are also available upon request.

5 Conclusions

In this paper, we examine whether bank lending in Russia increases prior to presidential elections. While the previous literature generally notes that state-owned banks can be used to influence political outcomes, we test a more specific hypothesis that both state-owned and privately held banks can be utilized to influence election outcomes in an electoral autocracy. Such regimes entail incentives and mechanisms that allow the authorities to exert political influence on the lending activities of both state-owned and private banks.

Our key finding is that Russian banks tend to boost their lending in the run-up to a presidential election. There are no significant differences between state-owned banks and private banks in this pattern. We next consider whether this lending surge is driven by demand or supply. We find that our results stand when controlling for the level of aggregate economic activity and that the lending surge results in deterioration of loan quality in the following year. This clearly suggests the pre-election lending surge is not related to exogenous economic events such as an economic expansion, leaving only political motivations as the potential driver.

Increased lending is observed for both household loans and corporate loans. Moreover, it is higher for larger banks and those banks most involved in lending activity. There is no evidence that banks with weak fundamentals or those with higher government deposits provide more loans than other banks. Our main conclusion, therefore, is that the increase in loan supply before elections is related to opaque forms of political interference in Russia's electoral autocracy.

Politicians in any democracy want to influence voters and surely see it as a cornerstone of the democratic electoral process. But Egorov and Sonin (2018) provide an excellent rationale for this phenomenon of political lending cycle by showing how elections can be used to signal the strength of a non-democratic leader. Elections outcomes can be extremely important to the party of power in a political system characterized as an electoral autocracy, a particularly apt description of Russia today. Our results on political lending cycles in Russia can therefore be viewed a contributing to our understanding of the interplay between politics and banking in non-democratic regimes. We show that the interplay between politics and banking in these regimes is not restricted to influencing state-owned banks but also includes influencing private banks. We further provide evidence that private banks are not immune to political interference. Further research, possibly concerning similar regimes (e.g. Turkey and Malaysia) is well-warranted to establish the exact mechanisms and economic structures that shape political lending cycles in non-democracies.

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Tables

Table 1 Descriptive statistics

Variable	Observations	Mean	Standard deviation
Total loan growth	131,116	0.015	0.122
Size	131,116	14.999	1.934
Capital/assets	131,116	0.208	0.141
Bad loans/loans	131,116	0.057	0.075
Loans/assets	131,116	0.626	0.175
Output index	131,116	0.633	9.050
State owner	131,116	0.052	0.221

Table 2 Main estimations: total loans

Dependent variable: total loan growth (mom)										
Election specification	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)		
	March	February	January	December	Jan–March	Dec–Feb	Oct–March	Sept–Feb		
Elections	0.009*** (0.003)	0.006** (0.003)	-0.003 (0.003)	0.023*** (0.003)	0.005*** (0.001)	0.009*** (0.002)	0.009*** (0.001)	0.010*** (0.001)		
State-owned	-0.000 (0.005)	0.000 (0.005)	-0.001 (0.005)	-0.000 (0.005)	-0.001 (0.005)	-0.001 (0.005)	-0.000 (0.005)	-0.000 (0.005)		
Elections×State-owned	-0.005	0.006	0.036***	-0.012	0.011**	0.010*	0.003	0.005		
	(0.008)	(0.008)	(0.012)	(0.012)	(0.005)	(0.006)	(0.004)	(0.004)		
Size	-0.003***	-0.003***	-0.003***	-0.003***	-0.003***	-0.003***	-0.003***	-0.003***		
	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)		
Capital/assets	0.038***	0.038***	0.038***	0.038***	0.038***	0.039***	0.039***	0.039***		
	(0.007)	(0.007)	(0.007)	(0.007)	(0.007)	(0.007)	(0.007)	(0.007)		
Bad loans/loans	-0.122***	-0.122***	-0.122***	-0.122***	-0.122***	-0.121***	-0.121***	-0.121***		
	(0.009)	(0.009)	(0.009)	(0.009)	(0.009)	(0.009)	(0.009)	(0.009)		
Loans/assets	-0.125***	-0.124***	-0.124***	-0.124***	-0.125***	-0.125***	-0.124***	-0.124***		
	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)		
Output index	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***		
	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)		
Constant	0.146***	0.145***	0.146***	0.142***	0.146***	0.144***	0.143***	0.143***		
	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)		
Observations	131,116	131,116	131,116	131,116	131,116	131,116	131,116	131,116		
Number of banks	1,209	1,209	1,209	1,209	1,209	1,209	1,209	1,209		
R2-within	0.0276	0.0275	0.0276	0.0280	0.0276	0.0278	0.0280	0.0280		
R2-between	0.0739	0.0740	0.0738	0.0734	0.0739	0.0737	0.0725	0.0722		
R2-overall	0.00513	0.00512	0.00513	0.00531	0.00513	0.00519	0.00532	0.00536		

Note: Monthly dummies included but not reported.

Table 3 Bad loan growth one year after elections

	(1)	(2)
Dependent variable	Bad loan growth	Overdue loan growth
	(year-on-year)	(year-on-year)
Elections	0.451***	0.395**
	(0.088)	(0.167)
Elections×State-owned	-0.109	-0.742*
	(0.248)	(0.435)
State-Owned	-0.345***	-0.070
	(0.128)	(0.413)
Size	0.016	0.011
	(0.021)	(0.046)
Capital/assets	-0.223	-0.548
•	(0.300)	(0.765)
Loans/assets	0.517**	1.900***
	(0.229)	(0.477)
Output index	0.166***	0.383***
•	(0.029)	(0.052)
Constant	0.403	0.531
	(0.348)	(0.826)
Observations	10,256	9,649
Number of banks	1,118	1,087
R2-within	0.0175	0.0142
R2-between	3.40e-05	0.0126
R2-overall	0.0115	0.00826

Note: lagged explanatory variables (12 months). Monthly dummies included but not reported.

Table 4 Main estimations: firm loans

Dependent variable: firm lo	oan growth (mom)							
Election specification	(1) March	(2) February	(3) January	(4) December	(5) Jan–March	(6) Dec–Feb	(7) Oct–March	(8) Sept–Feb
Elections	0.011*** (0.003)	0.007*** (0.003)	-0.000 (0.003)	0.002 (0.003)	0.006*** (0.002)	0.003* (0.002)	0.005*** (0.001)	0.004*** (0.001)
State-owned	-0.003 (0.005)	-0.003 (0.005)	-0.003 (0.005)	-0.003 (0.005)	-0.003 (0.005)	-0.003 (0.005)	-0.003 (0.005)	-0.003 (0.005)
Elections×State-owned	-0.003 (0.011)	-0.003 (0.007)	0.010 (0.007)	0.005 (0.017)	0.001 (0.005)	0.004 (0.006)	0.006 (0.004)	0.007 (0.004)
Size	-0.003*** (0.001)							
Capital/assets	-0.008 (0.008)	-0.008 (0.008)	-0.009 (0.008)	-0.008 (0.008)	-0.008 (0.008)	-0.008 (0.008)	-0.008 (0.008)	-0.008 (0.008)
Bad loans/loans	-0.160*** (0.011)							
Loans/assets	-0.058*** (0.005)	-0.058*** (0.005)	-0.057*** (0.005)	-0.058*** (0.005)	-0.058*** (0.005)	-0.058*** (0.005)	-0.058*** (0.005)	-0.058*** (0.005)
Output index	0.001*** (0.000)							
Constant	0.069*** (0.012)	0.069*** (0.012)	0.069*** (0.012)	0.069*** (0.012)	0.069*** (0.012)	0.069*** (0.012)	0.068*** (0.012)	0.069*** (0.012)
Observations	129,616	129,616	129,616	129,616	129,616	129,616	129,616	129,616
Number of banks	1,194	1,194	1,194	1,194	1,194	1,194	1,194	1,194
R2-within	0.0165	0.0164	0.0163	0.0164	0.0165	0.0164	0.0165	0.0165
R2-between	0.0212	0.0213	0.0210	0.0214	0.0212	0.0213	0.0215	0.0215
R2-overall	0.00805	0.00799	0.00798	0.00792	0.00804	0.00794	0.00805	0.00799

Note: Monthly dummies included but not reported.

Table 5 Main estimations: household loans

Dependent variable: househ	nold loan growth (m	om)						
Election specification	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	March	February	January	December	Jan–March	Dec–Feb	Oct–March	Sept–Feb
Elections	0.011***	0.011***	0.003	0.008*	0.008***	0.007***	0.007***	0.007***
	(0.003)	(0.003)	(0.003)	(0.004)	(0.002)	(0.002)	(0.002)	(0.002)
State-owned	0.003	0.003	0.003	0.003	0.003	0.003	0.003	0.004
	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.005)	(0.005)
Elections×State-owned	0.003 (0.009)	-0.003 (0.007)	0.002 (0.012)	-0.015 (0.015)	0.000 (0.006)	-0.006 (0.008)	-0.004 (0.008)	-0.006 (0.008)
Size	-0.010***	-0.010***	-0.010***	-0.010***	-0.010***	-0.010***	-0.010***	-0.010***
	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)
Capital/assets	-0.075***	-0.075***	-0.075***	-0.075***	-0.075***	-0.075***	-0.075***	-0.075***
	(0.008)	(0.008)	(0.008)	(0.008)	(0.008)	(0.008)	(0.008)	(0.008)
Bad loans/loans	-0.169***	-0.169***	-0.170***	-0.170***	-0.169***	-0.169***	-0.169***	-0.169***
	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)
Loans/assets	-0.021***	-0.021***	-0.021***	-0.021***	-0.021***	-0.021***	-0.021***	-0.021***
	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)
Output index	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***
	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)
Constant	0.190***	0.190***	0.190***	0.189***	0.189***	0.189***	0.189***	0.189***
	(0.012)	(0.012)	(0.012)	(0.012)	(0.012)	(0.012)	(0.012)	(0.012)
Observations	128,350	128,350	128,350	128,350	128,350	128,350	128,350	128,350
Number of banks	1,170	1,170	1,170	1,170	1,170	1,170	1,170	1,170
R2-within	0.0179	0.0179	0.0178	0.0178	0.0180	0.0179	0.0180	0.0180
R2-between	0.00351	0.00359	0.00365	0.00364	0.00357	0.00360	0.00359	0.00356
R2-overall	0.00731	0.00724	0.00715	0.00717	0.00731	0.00721	0.00729	0.00728

Note: Monthly dummies included but not reported.

Table 6 Lending growth of large and small banks

Dependent variable: total loan growth (mom)											
Large banks	(1) March	(2) February	(3) January	(4) December	(5) Jan–March	(6) Dec–Feb	(7) Oct–March	(8) Sept–Feb			
Elections	0.011*** (0.003)	0.008*** (0.003)	-0.004 (0.004)	0.025*** (0.004)	0.006*** (0.002)	0.009*** (0.002)	0.010*** (0.001)	0.010*** (0.001)			
Elections×State-owned	-0.016* (0.008)	0.004 (0.008)	0.043*** (0.014)	-0.017 (0.013)	0.008 (0.005)	0.010 (0.007)	0.003 (0.004)	0.006 (0.005)			
Observations	69,729	69,729	69,729	69,729	69,729	69,729	69,729	69,729			
Number of banks	803	803	803	803	803	803	803	803			
R2-within	0.0398	0.0397	0.0398	0.0403	0.0398	0.0400	0.0403	0.0404			
R2-between	0.00939	0.00954	0.00939	0.00888	0.00956	0.00942	0.00905	0.00906			
R2-overall	0.0104	0.0104	0.0104	0.0106	0.0104	0.0105	0.0106	0.0106			

Dependent variable: total loan growth (mom)										
Small banks	(1) March	(2) February	(3) January	(4) December	(5) Jan–March	(6) Dec–Feb	(7) Oct–March	(8) Sept–Feb		
Elections	0.006 (0.004)	0.004 (0.004)	-0.001 (0.005)	0.023*** (0.006)	0.003 (0.002)	0.008*** (0.003)	0.008*** (0.002)	0.010*** (0.002)		
Elections×State-owned	0.048*** (0.011)	0.024 (0.020)	0.026 (0.026)	-0.009 (0.043)	0.034*** (0.013)	0.015 (0.009)	0.013* (0.007)	0.006 (0.007)		
Observations	61,387	61,387	61,387	61,387	61,387	61,387	61,387	61,387		
Number of banks	815	815	815	815	815	815	815	815		
R2-within	0.0268	0.0268	0.0267	0.0271	0.0268	0.0269	0.0271	0.0272		
R2-between	0.0469	0.0468	0.0471	0.0474	0.0465	0.0461	0.0454	0.0454		
R2-overall	0.00431	0.00428	0.00427	0.00445	0.00431	0.00435	0.00443	0.00448		

Note: Bank-level controls, output variable and monthly dummies included but not reported.

Table 7 Lending growth by involvement in lending

Dependent variable: total loan growth (mom)											
High lending banks	(1) March	(2) February	(3) January	(4) December	(5) Jan–March	(6) Dec–Feb	(7) Oct–March	(8) Sept–Feb			
Elections	0.006** (0.003)	0.007*** (0.003)	-0.004 (0.003)	0.026*** (0.004)	0.003** (0.001)	0.009*** (0.002)	0.010*** (0.001)	0.011*** (0.001)			
Elections×State-owned	0.008 (0.006)	-0.006 (0.007)	0.018** (0.008)	-0.005 (0.010)	0.006* (0.003)	0.002 (0.004)	0.001 (0.004)	0.001 (0.004)			
Observations	68,439	68,439	68,439	68,439	68,439	68,439	68,439	68,439			
Number of banks	1,090	1,090	1,090	1,090	1,090	1,090	1,090	1,090			
R2-within	0.113	0.113	0.113	0.114	0.113	0.113	0.114	0.114			
R2-between	0.0513	0.0515	0.0513	0.0514	0.0515	0.0518	0.0523	0.0528			
R2-overall	0.0845	0.0845	0.0845	0.0852	0.0845	0.0848	0.0853	0.0854			

Dependent variable: total loan growth (mom)										
Low lending banks	(1) March	(2) February	(3) January	(4) December	(5) Jan–March	(6) Dec–Feb	(7) Oct–March	(8) Sept–Feb		
Elections	0.010** (0.005)	0.002 (0.005)	0.000 (0.005)	0.024*** (0.006)	0.004 (0.003)	0.008*** (0.003)	0.008*** (0.002)	0.009*** (0.002)		
Elections×State-owned	-0.022 (0.015)	0.032** (0.016)	0.068** (0.028)	-0.031 (0.025)	0.024** (0.011)	0.025* (0.014)	0.008 (0.008)	0.011 (0.010)		
Observations	62,677	62,677	62,677	62,677	62,677	62,677	62,677	62,677		
Number of banks	1,131	1,131	1,131	1,131	1,131	1,131	1,131	1,131		
R2-within	0.0541	0.0541	0.0542	0.0544	0.0542	0.0543	0.0544	0.0544		
R2-between	0.00273	0.00278	0.00279	0.00262	0.00283	0.00282	0.00268	0.00267		
R2-overall	0.0190	0.0190	0.0190	0.0192	0.0190	0.0191	0.0191	0.0192		

Note: Bank-level controls, output variable and monthly dummies included but not reported.

Table 8 Estimations accounting for government deposits

Dependent variable: total loan growth (mom)										
Election specification	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)		
	March	February	January	December	Jan–March	Dec–Feb	Oct–March	Sept–Feb		
Elections	0.0167***	0.00922***	-0.00319	0.0242***	0.00749***	0.0101***	0.0122***	0.0124***		
	(0.00312)	(0.00328)	(0.00318)	(0.00352)	(0.00171)	(0.00169)	(0.00124)	(0.00120)		
Gov deposits	0.00485***	0.00487***	0.00452**	0.00496***	0.00491***	0.00502***	0.00535***	0.00527***		
	(0.00179)	(0.00179)	(0.00179)	(0.00178)	(0.00181)	(0.00179)	(0.00183)	(0.00183)		
Elections×Gov deposits	-0.00160 (0.00617)	-0.00807 (0.00654)	0.00893 (0.00742)	-0.00461 (0.00857)	-0.000232 (0.00362)	-0.00139 (0.00370)	0.000464 (0.00297)	0.00110 (0.00317)		
State-owned	0.00262	0.00249	0.00186	0.00270	0.00175	0.00184	0.00178	0.00161		
	(0.00523)	(0.00523)	(0.00532)	(0.00518)	(0.00531)	(0.00528)	(0.00530)	(0.00529)		
Elections×State-owned	-0.00617 (0.00832)	0.00222 (0.00818)	0.0348*** (0.0126)	-0.0119 (0.0126)	0.0108** (0.00527)	0.00880 (0.00633)	0.00230 (0.00433)	0.00325 (0.00434)		
Size	-0.00115	-0.00119	-0.00125	-0.00105	-0.00110	-0.00104	-0.000726	-0.000690		
	(0.000764)	(0.000763)	(0.000763)	(0.000758)	(0.000764)	(0.000762)	(0.000760)	(0.000760)		
Capital/assets	0.0478***	0.0476***	0.0470***	0.0479***	0.0480***	0.0484***	0.0499***	0.0500***		
	(0.00904)	(0.00905)	(0.00902)	(0.00901)	(0.00906)	(0.00905)	(0.00905)	(0.00905)		
Bad loans/loans	-0.102***	-0.103***	-0.103***	-0.102***	-0.102***	-0.102***	-0.100***	-0.101***		
	(0.0105)	(0.0105)	(0.0105)	(0.0105)	(0.0105)	(0.0105)	(0.0105)	(0.0105)		
Loans/assets	-0.125***	-0.125***	-0.124***	-0.124***	-0.125***	-0.125***	-0.125***	-0.125***		
	(0.00682)	(0.00682)	(0.00681)	(0.00681)	(0.00682)	(0.00682)	(0.00682)	(0.00681)		
Output index	0.00104***	0.000944***	0.00104***	0.000985***	0.000988***	0.000948***	0.000972***	0.000934***		
	(0.000172)	(0.000177)	(0.000171)	(0.000172)	(0.000173)	(0.000174)	(0.000172)	(0.000173)		
Constant	0.111***	0.112***	0.113***	0.104***	0.110***	0.107***	0.101***	0.100***		
	(0.0132)	(0.0132)	(0.0132)	(0.0131)	(0.0132)	(0.0132)	(0.0131)	(0.0132)		
Observations	101,562	101,562	101,562	101,562	101,562	101,562	101,562	101,562		
Number of banks	1,117	1,117	1,117	1,117	1,117	1,117	1,117	1,117		
R2-within	0.0259	0.0257	0.0257	0.0261	0.0258	0.0259	0.0265	0.0265		
R2-between	0.0599	0.0601	0.0598	0.0607	0.0600	0.0604	0.0588	0.0587		
R2-overall	0.0399	0.00488	0.0398	0.00522	0.00498	0.00506	0.00548	0.0387		

Table 9 Estimations accounting for weak fundamentals: interaction with capital-to-assets ratio

Dependent variable: total loan growth (mom)										
Election specification	(1) March	(2) February	(3) January	(4) December	(5) Jan–March	(6) Dec–Feb	(7) Oct–March	(8) Sept–Feb		
Elections	0.012*** (0.005)	0.010** (0.004)	-0.009* (0.005)	0.024*** (0.005)	0.005* (0.003)	0.008*** (0.003)	0.008*** (0.002)	0.008*** (0.002)		
State-owned	-0.000 (0.005)	-0.000 (0.005)	-0.001 (0.005)	-0.000 (0.005)	-0.001 (0.005)	-0.001 (0.005)	-0.001 (0.005)	-0.001 (0.005)		
Capital/assets	0.039*** (0.007)	0.038*** (0.007)								
Elections×State-owned	-0.006 (0.008)	0.005 (0.008)	0.038*** (0.012)	-0.013 (0.012)	0.011** (0.005)	0.010* (0.006)	0.004 (0.004)	0.005 (0.004)		
Elections×Capital/assets	-0.013 (0.021)	-0.018 (0.021)	0.030 (0.023)	-0.006 (0.025)	-0.002 (0.013)	0.002 (0.012)	0.005 (0.009)	0.010 (0.008)		
Bad loans/loans	-0.122*** (0.009)	-0.121*** (0.009)	-0.122*** (0.009)	-0.122*** (0.009)	-0.121*** (0.009)	-0.121*** (0.009)	-0.121*** (0.009)	-0.121*** (0.009)		
Size	-0.003*** (0.001)									
Loans/assets	-0.125*** (0.006)	-0.124*** (0.006)	-0.124*** (0.006)	-0.124*** (0.006)	-0.124*** (0.006)	-0.125*** (0.006)	-0.124*** (0.006)	-0.124*** (0.006)		
Output index	0.001*** (0.000)									
Constant	0.146*** (0.011)	0.145*** (0.011)	0.146*** (0.011)	0.142*** (0.011)	0.146*** (0.011)	0.144*** (0.011)	0.144*** (0.011)	0.144*** (0.011)		
Observations	131,116	131,116	131,116	131,116	131,116	131,116	131,116	131,116		
Number of banks	1,209	1,209	1,209	1,209	1,209	1,209	1,209	1,209		
R2-within	0.0276	0.0275	0.0276	0.0280	0.0276	0.0278	0.0280	0.0281		
R2-between	0.0739	0.0740	0.0738	0.0734	0.0738	0.0737	0.0723	0.0719		
R2-overall	0.00514	0.00514	0.00513	0.00530	0.00513	0.00519	0.00532	0.00535		

Note: Monthly dummies included but not reported.

Table 10 Estimations accounting for weak fundamentals: interaction with bad loans ratio

Dependent variable: total loan growth (mom)								
Election specification	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	March	February	January	December	Jan–March	Dec–Feb	Oct–March	Sept–Feb
Elections	0.012***	0.008**	-0.010***	0.029***	0.004**	0.009***	0.010***	0.010***
	(0.003)	(0.003)	(0.004)	(0.004)	(0.002)	(0.002)	(0.001)	(0.001)
State-owned	-0.000 (0.005)	-0.000 (0.005)	-0.001 (0.005)	-0.000 (0.005)	-0.001 (0.005)	-0.001 (0.005)	-0.001 (0.005)	-0.001 (0.005)
Bad loans/loans	-0.121***	-0.121***	-0.124***	-0.120***	-0.122***	-0.121***	-0.119***	-0.120***
	(0.009)	(0.009)	(0.009)	(0.009)	(0.009)	(0.009)	(0.010)	(0.010)
Elections×State-owned	-0.005 (0.008)	0.006 (0.008)	0.034*** (0.012)	-0.011 (0.012)	0.011** (0.005)	0.010* (0.006)	0.004 (0.004)	0.005 (0.004)
Elections×Bad loans/loans	-0.042 (0.043)	-0.037 (0.037)	0.119*** (0.038)	-0.108** (0.044)	0.009 (0.024)	-0.008 (0.023)	-0.015 (0.017)	-0.008 (0.015)
Size	-0.003***	-0.003***	-0.003***	-0.003***	-0.003***	-0.003***	-0.003***	-0.003***
	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)	(0.001)
Capital/assets	0.038***	0.038***	0.038***	0.038***	0.038***	0.038***	0.038***	0.038***
	(0.007)	(0.007)	(0.007)	(0.007)	(0.007)	(0.007)	(0.007)	(0.007)
Loans/assets	-0.125***	-0.124***	-0.124***	-0.124***	-0.124***	-0.124***	-0.124***	-0.124***
	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)	(0.006)
Output index	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***	0.001***
	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)	(0.000)
Constant	0.146***	0.145***	0.146***	0.141***	0.146***	0.144***	0.143***	0.143***
	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)	(0.011)
Observations	131,116	131,116	131,116	131,116	131,116	131,116	131,116	131,116
Number of banks	1,209	1,209	1,209	1,209	1,209	1,209	1,209	1,209
R2-within	0.0276	0.0276	0.0277	0.0280	0.0276	0.0278	0.0279	0.0280
R2-between	0.0739	0.0737	0.0744	0.0725	0.0739	0.0735	0.0717	0.0718
R2-overall	0.00514	0.00512	0.00518	0.00536	0.00513	0.00520	0.00534	0.00537

Note: Monthly dummies included but not reported.

Table 11 Excluding one election episode (12 months before and after the election) at a time

Dependent variable: total loan	growth (mom)							
Election specification	(1) March	(2) February	(3) January	(4) December	(5) Jan–March	(6) Dec–Feb	(7) Oct–March	(8) Sept–Feb
			Excludi	ng 2004				
Elections	0.010*** (0.003)	0.004 (0.003)	-0.003 (0.003)	0.022*** (0.003)	0.004** (0.002)	0.008*** (0.002)	0.008*** (0.001)	0.009*** (0.001)
Elections×State-owned	-0.006 (0.008)	0.001 (0.008)	0.037*** (0.012)	-0.012 (0.012)	0.011** (0.005)	0.009 (0.006)	0.003 (0.004)	0.004 (0.004)
			Excludi	ng 2008				
Elections	0.009*** (0.003)	0.006* (0.003)	-0.012*** (0.004)	0.022*** (0.004)	0.002 (0.002)	0.006*** (0.002)	0.005*** (0.001)	0.005*** (0.001)
Elections×State-owned	-0.002 (0.009)	0.011 (0.010)	0.032** (0.013)	-0.017 (0.016)	0.012* (0.007)	0.009 (0.007)	0.004 (0.005)	0.005 (0.005)
			Excludi	ng 2012				
Elections	0.013*** (0.003)	0.009** (0.003)	0.005 (0.004)	0.021*** (0.004)	0.009*** (0.002)	0.011*** (0.002)	0.012*** (0.001)	0.012*** (0.001)
Elections×State-owned	-0.002 (0.010)	0.014 (0.011)	0.038** (0.017)	-0.002 (0.014)	0.015** (0.006)	0.017** (0.008)	0.009* (0.005)	0.007 (0.006)
			Excludi	ng 2018				
Elections	0.010*** (0.003)	0.008*** (0.003)	-0.006* (0.003)	0.026*** (0.004)	0.005*** (0.002)	0.009*** (0.002)	0.010*** (0.001)	0.011*** (0.001)
Elections×State-owned	-0.010 (0.009)	-0.003 (0.008)	0.039*** (0.015)	-0.018 (0.016)	0.007 (0.004)	0.006 (0.007)	-0.001 (0.004)	0.002 (0.005)

Note: Bank-level controls, output variable and monthly dummies included but not reported.

Table 12 Placebo elections defined for March 2006, 2010 and 2014

Dependent variable: total loan growth (mom)								
Election specification	(1) March	(2) February	(3) January	(4) December	(5) Jan–March	(6) Dec–Feb	(7) Oct–March	(8) Sept–Feb
Elections	-0.006** (0.003)	-0.007*** (0.003)	-0.001 (0.003)	-0.006* (0.003)	-0.005*** (0.001)	-0.005*** (0.002)	-0.002** (0.001)	-0.002* (0.001)
State-owned	-0.001 (0.006)	-0.001 (0.006)	-0.001 (0.006)	-0.001 (0.006)	-0.001 (0.006)	-0.001 (0.006)	-0.001 (0.005)	-0.001 (0.006)
Elections×State-owned	-0.003 (0.009)	0.015* (0.008)	0.001 (0.012)	-0.007 (0.011)	0.005 (0.006)	0.003 (0.005)	0.003 (0.005)	0.004 (0.004)
Size	-0.003*** (0.001)	-0.003*** (0.001)	-0.003*** (0.001)	-0.003*** (0.001)	-0.003*** (0.001)	-0.003*** (0.001)	-0.003*** (0.001)	-0.003*** (0.001)
Capital/assets	0.036*** (0.007)							
Bad loans/loans	-0.125*** (0.010)							
Loans/assets	-0.128*** (0.006)							
Output index	0.001*** (0.000)							
Constant	0.148*** (0.011)	0.148*** (0.011)	0.148*** (0.011)	0.150*** (0.011)	0.149*** (0.011)	0.150*** (0.011)	0.149*** (0.011)	0.149*** (0.011)
Observations	126,466	126,466	126,466	126,466	126,466	126,466	126,466	126,466
Number of banks	1,205	1,205	1,205	1,205	1,205	1,205	1,205	1,205
R2-within	0.0282	0.0282	0.0281	0.0282	0.0282	0.0282	0.0282	0.0282
R2-between	0.0755	0.0755	0.0755	0.0756	0.0756	0.0758	0.0756	0.0756
R2-overall	0.00529	0.00530	0.00528	0.00529	0.00530	0.00530	0.00528	0.00528

Note: Monthly dummies included but not reported.

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