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Welfare and food security response of animal feed and water resource scarcity in Northern Ethiopia

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Abstract

The scarcity of grazing and water for an animal has a negative effect on household welfare and food security either by affecting livestock production directly, affecting crop or off-farm income due to labor reallocation or through its direct impact on time leisure consumption.

The economic impacts of resource (grazing and water) scarcity on welfare are undermined. Thus, a better understanding that is derived from the factual evidence is required. The first objective of this paper is to explore the link between natural resource scarcity and per capita food consumption expenditure (PCFE) as proxy for welfare and food security followed by the second objective of analyzing whether this effect is uniform across all quantile groups and there is gender differential effect using distance and shadow price as resource scarcity indicators. The paper used a relatively unique data set from a randomly drawn 518 sample farmers in Northern Ethiopia. To address our first objective, we employ the IV two-stage least square estimation for welfare and probit model for food security drawing on non-separable farm household model.

Our estimates show that about 48% of the households were food secure while 52% were food insecure. Our results confirmed the theoretical prediction that resource scarcity affects household PCFE and food security adversely as predicted by the downward spiral hypothesis. The results indicate that animal feed and water scarcity have an important impact on welfare and food security. As expected, in aggregate, reducing time spent searching for water per day leads to an increase in PCFE and food security. Similarly, a decrease in time wastage for searching grazing increase PCFE and food security respectively, and an increment of PCFE and food security is achieved by a reduction in crop residue transporting time per day.

The gender differential analysis signals that increasing resource scarcity results in low PCFE and food security, with the male are considerably likely to have less food consumption expenditure and being food insecure more as compared to female households. The total impact of time spent searching for water, grazing, and transporting straw on per PCFE is -0.142% , -0.102% , and -0.092% , respectively, and decreasing reaching time to a water, grazing, and straw source by 0.6 min will increase PCFE by 354 ETB, 254 ETB, and 229 ETB for the median household. Depending on results from the quantile regression, the effect of water and feed scarcity is not uniform across the food consumption distribution.

Keywords: Food consumption, Resource scarcity: welfare, Food security: IV, Probit, Ethiopia

Introduction

Many studies have established that the rural poor in developing countries are heavily dependent on local natural resources for their subsistence (e.g., Narain et al. 2008) and that the depletion of these resources hurts the poor more (Khan 2008). Land degradation in Sub-Saharan Africa (SSA) remains a substantial problem to spur rural poverty (Bhattacharya and Innes 2006; Tesfa and Mekuriaw 2014). It directly aggravates poverty, by reducing the availability of environmental goods and services and by increasing the labor input needed to seek for them (Bezabih and Berhane 2014; Lal and Stewart 2010). The critical shortage of those resources has negative implications for agricultural production and food security, particularly for poor people who rely on agriculture as a source of food and spend considerable time to collecting these resources (Mekonnen et al. 2015; Yilma et al. 2011).

Poverty and resource degradation appear to go hand in hand in SSA. Resource degradation by all accounts is rampant in the region. In East Africa, livestock production depends on quantity and quality of grazing feed and water. About 10% of cropland is used to produce crop residues for feeding livestock, and animals in the extensive system need more water per animal (Bezabih and Berhane 2014). Increasing scarcity of grazing and water for an animal can be a significant burden to households, as grazing and water are a key factor in livestock production. Thus, the scarcity of these resources may directly impact agriculture or indirectly by reallocating factors of production, namely labor from agriculture, food preparation, and leisure activities to searching and collecting these scarce resources. Reductions in agricultural output stemming from less labor input are very likely to have detrimental welfare and food security consequence (Cooke 1998; Kumar and Hotchkiss 1988; Mekonnen et al. 2015). The downward spiral hypothesis states that people in poverty are forced to deplete resources to survive, and this environmental depletion further impoverishes them (Ostrom et al. 1999).

In many studies of Africa, grazing land and water scarcity are frequently mentioned constraints for animal farming activities (Bezabih and Berhane 2014; Tegegne 2012). Grazing and water scarcity may be less problematic in developed countries where there are available substitutes but can have a huge impact on household welfare in developing countries like Ethiopia. Resource depletion, in the country, has contributed to the existing problem of food insecurity and is still a real threat to the agricultural farming (Bewket 2011). World Bank (2012) reported that the cost of environmental degradation is almost 8% of GDP across countries consisting 40% of the developing countries. More specifically, environmental depletion, in the study area, has reached a critical stage which poses a major threat to the agriculture production and welfare (Gebregziabher et al. 2008). Households with scarcity may walk longer distances to search and collect these resources, thereby leaving less labor for leisure, food production, and preparation (Bezabih and Berhane 2014; Cooke et al. 2008).

The literature suggests that as a result of increasing resource scarcity, many households increase the time they spend on collecting them. Overall, the scarcity has negative implications for agricultural production and the food security by diminishing households' food supply and incomes, and hence their capacity to achieve food and nutrition (Cooke et al. 2008; Damte et al. 2012; Mekonnen et al. 2015; Tangka and Jabbar 2005). The findings of Cooke (1998) revealed that most of the reallocated time for searching and collecting the scarce resources come from leisure before agricultural labor time is reduced. One early

analysis conducted by Bandyopadhyay et al. (2011) also indicates that amount of biomass negatively affected rural per capita consumption expenditure in Malawi. Grazing and water scarcity in Ethiopia can affect household welfare in different ways.

Poor farmers may not have access to alternative feed resources and may increase the time spent on searching grazing, water and straw collection, reducing time on farming activities, food preparation, and leisure or household care. Thus, under situations where markets are imperfect, increasing resource scarcity can force households to reallocate labor and thereby reduce welfare. The degree to which labor allocated to an animal feeding and watering takes labor away from agricultural production likely depends on who in the household is engaged in farming when feed collection and grazing takes place (Arnold et al. 2003). It is commonly perceived that children and women are mostly responsible for the collection of feeding and watering an animal and that their scarcity increases the burden on these household members. Individual member's increased labor burden further reduces their overall human welfare (Mekonnen et al. 2015; Tangka and Jabbar 2005).

While the above studies estimate the effect of resource scarcity on time allocation and time reduction for farming, no study we are aware of examining the economic effect of grazing and water scarcity on welfare, which is ultimately what policymakers seek to know (Cooke et al. 2008; Khan 2008; Tangka and Jabbar 2005). This is of potentially relevant but less well studied. In this study, we are able to estimate the effect of grazing, water, and straw scarcity on per capita food consumption expenditure (PCFE) as a proxy for welfare and food security using distance and shadow price¹ as a proxy for scarcity indicator of these resources by exploiting household survey from Northern Ethiopia. Our analysis is organized around five questions. First, what is the effect of this resource scarcity on welfare (PCFE)? Second, how does resource scarcity affect household food security? Third, is this effect uniform across all quantile groups? Fourth, is there gender differential effect? And, fifth, what is the effect of the scarcity on the total welfare?

In line to this, we hypothesize that the scarcity has a negative effect on households' food security and welfare (PCFE) either by affecting livestock production directly, affecting crop, and off-farm income or through its direct impact on time for leisure consumption drawing on a non-separable farm household model. We also hypothesize that the effect of these scarce resources is not uniform across the food consumption distribution. In aggregate, the principal findings confirmed the theoretical prediction that resource scarcity affects household welfare (PCFE) and food security adversely as predicted by the downward spiral hypothesis. The estimated result from both distance and shadow price revealed that reducing time spent for searching the water, grazing, and collecting crop residue leads to an increase in welfare (PCFE) and food security.

This paper builds on the existing literature in a number of respects. In this paper, we contribute to the literature by using a unique dataset to investigate how the distance to or the shadow price of water, grazing, and crop residue affects PCFE and food security. We are able to estimate causal relationships with our data because, unlike previous studies, we collected information on the entire set of consumption expenditure, along with the distance to grazing, water and crop residue of each household. Furthermore, unlike the previous studies, we use distance and shadow price as a proxy measure of resource scarcity. This paper joins the relatively scarce, empirical literature on this topic in Africa, one that is dominated by South Asian cases such as Nepal.²

Literature review

Review of background

The contribution of livestock to the world's food supply, family nutrition, incomes, employment, soil fertility, and transport helps for the reduction of food insecurity and poverty (Randolph et al. 2007). Livestock also gives a safety net in the form of liquid assets and a strategy of food production diversification (Freeman et al. 2007). In Ethiopia, the agricultural sector is a cornerstone of the economic and social life of the people. Livestock contribution accounts for 40% of total agricultural GDP, excluding the values of draft power, manure, and transport service (Asresie and Zemedu 2015). Despite its large population size, the contribution of livestock production to agriculture is deteriorating (Ilyin 2011). Livestock production in Ethiopia depends on the quantity of grazing feed and water (Bezabih and Berhane 2014). This sector is a key player in increasing water use and water depletion (Steinfeld et al. 2006).

Ethiopia is a home of 35 million tropical livestock unit (TLU), and on average, one TLU requires about 25 l of water per day and the total daily water requirement for livestock is estimated at 875 million liters amounting to about 320 billion liters per year. Total grazing land in Tigray is estimated to be 47,431 km² while tropical livestock unit (TLU) per km² of grazing land was increased from 44,000 TLU in 2001/2002 to 55,000 TLU in 2007/2008. Thus, TLU per km² grazing land in the region is above half for each year due to greater population density, larger herd sizes, and relatively fixed grazing land resources (Tilahun and Schmidt 2012). Both human and livestock suffer from its shortage. Most of the year, animals have to walk long distances in search of water. Major feed resources, in Ethiopia, are crop residues and natural pasture, but their availability is gradually declining as a result of crop expansion, settlement, and land degradation (Gebremedhin 2009; Yimer 2005).

The case study by Belay et al. (2013) indicated that the most important problems of livestock production perceived were feed shortage (100%) and water shortage (27%) during the dry season in Ethiopia. Feed and water deficits start in December, when the natural pastures are at their lowest quantity and the supply of stored crop residues is starting to diminish (Sileshi et al. 2003). In the study region, Gebremedhin (2009) and Yimer (2005) also revealed that natural grazing is diminishing over time due to the high degree of degradation, resulting in high TLU per km² of grazing land. The estimated crop residues are found to be about 1,229,651 tons dry matter/year while the region has an estimated 878,322 ha of arable land available for crop production, contributing about 45% of the animal feed demand. Felleke and Geda (2001) stated that 73% of the feed is provided from natural grazing, 14% from crop residues, and the remaining 13% from other feed sources. A recent study of Bishu (2014) in Tigray indicated that there is water shortage for livestock drinking (34%) and feed shortages (7%). There is also a shortage of labor for livestock management (Tegegne 2012).

Empirical literature

Dasgupta (2007) warns that the average per capita consumption level may decline with degradation of resources. Aggrey et al. (2010) showed that shortage of firewood and fuelwood were positively linked with poverty in Uganda. The findings of Khan (2008)

in Pakistan supported that environmental degradation hurts the poor more. The study of Kumar and Hotchkiss (1988) suggested that shortage of firewood has adverse effects on agricultural production, food consumption, and nutrition in Nepal. Poor farmers, who are directly dependent on these local natural resources, are highly affected by the resource scarcity. Cooke's (1998) result revealed that a reallocation of time away from leisure occurred as environmental goods become scarcer in Nepal. In addition, the findings of Tangka and Jabbar (2005) in Kenya shows that feed scarcity increases livestock traveling distances in search of feed and water that increase household time for collection, resulting in lower livestock and crop output which further diminishes households' food and nutrition security.

Likewise, Cooke et al. (2008) found a negative effect of resource scarcity on health, labor burden, and agriculture in Nepal. Bhattacharya and Innes (2006) highlighted that forest degradation spurs rural poverty in Sub-Saharan Africa. According to the study of Bandyopadhyay et al. (2011) in Malawi, more time spent on scarce fuelwood collection was associated with negative welfare. Baland et al. (2010) in their result indicate that an increase in firewood collection time lowers living standards of households in Nepal. The study of Aluko (2004) showed that deterioration in the quality of life increases with increasing environmental degradation in Niger. Mekonnen et al. (2015), in their analysis, show that fuelwood scarcity has a negative impact on time spent on agriculture; however, scarcity of water had no effect on time spent on agriculture in Ethiopia. Likewise, Mekonnen et al. (2017) on their analysis indicated that agricultural productivity decreases with increasing time spent on collecting animal dung but increases with time spent on collecting crop residue. The paper by Boone et al. (2011) suggests that long distance to water source increase water gathering time in Madagascar.

In spite of the recognized contributions of the existing studies, none of the above studies examine the effect of grazing and water on welfare and food security (Tangka and Jabbar, 2005; Cooke et al. 2008). Therefore, this study makes a noteworthy contribution in pointing out the relevance of improving feed and water management for the animal.

Theoretical model

The contribution of livestock to food and nutritional security in developing countries is significant (Swanepoel et al. 2010). In a mixed crop-livestock, Ethiopia owns a significantly large livestock population and its production mainly depends on natural resources such as grazing land, water, and own crop residue (Bezabih and Berhane 2014). Ethiopian farmers usually experience a very serious seasonal fluctuation in fodder and water availability for the animal. The dependence on these resources implies that scarcity can have a huge impact on household welfare (Bewket 2011; Bezabih and Berhane 2014). In rural farm households, where the farmer is engaged in both crop and livestock production activities, total time endowment is divided into three main activities: farm activities, off-farm activities, and leisure. However, considering the scarcity of these resources, the total time endowment will further include the 4th, collecting scarce resource activities.

We start with the downward spiral hypothesis which states that people in poverty are forced to deplete resources to survive, and this environmental depletion further

impoverishes them (Ostrom et al. 1999). It is supposed that the scarcity of resource can affect household well-being either by affecting livestock production directly, affecting crop and off-farm income (via labor reallocation), or through its direct impact on time for food preparation or leisure consumption (Cooke et al. 2008; Mekonnen et al. 2015). To conceptualize the effect of resource scarcity on welfare and food security, we develop a theoretical model within the framework of household utility model following the work of Strauss (1986a) and later used by Faridi and Wadood (2010). We suppose that rural households are characterized as both producers and consumers of their food, and thus, household strictly quasi-concave utility function based on the framework of consumer demand and production theories is presented as follows:

$$U_i = U(C_i, C_n, C_m, L_i; \Gamma) \tag{1}$$

where U_i is a utility function that is twice differentiable, increasing in its arguments, and strictly quasi-concave; C_i and C_n are a vector of home produced food and non-food goods consumed by the i th household; C_m is a market-purchased goods consumed; and L_i is leisure and Γ is the vector of household socio-demographic variables. C_n in this case represents the demand for non-food items such as education, health, and housing. Equation (1) leads us to the generalized utility function developed by Becker's (1981), which requires that production decision is first made to maximize profit and household maximizes utility using this maximum profit consecutively (Strauss 1986a). The meal production is a function of agricultural goods (Q_i), off-farm income (E), and fuel sources such as straw or dung (E_f) as well as labor days the household spend on searching grazing land, water, and crop residue (L_c). The production of household goods is also influenced by the vector of household characteristics.

$$C_i = C(E_f, Q_i, E, L_c; \Xi) \tag{2}$$

The rural household is assumed to maximize its utility subject to farm production, income, and time constraints specified as:

$$F(Q_i, L, L_c, K, A) = 0 \tag{3}$$

Equation (3) is a typical household implicit production function for food, Q_i produced at home and assumed to be twice differentiable, increasing in outputs, decreasing in inputs, and strictly convex; L is total labor used on the farm; L_c is the time spent on searching grazing, water, and collecting crop residue; K is the fixed capital stock and A is the farm size and the labor time is an important resource, denoted by T , and it is allocated among crop farming activities L_a , searching and collecting scarce resource L_c and leisure L_i :

$$T = L_a + L_c + L_i \tag{4}$$

At the same time, the income constraint for the rural household is given by

$$P_i(Q_i - C_i) - P_n C_n - P_m C_m - W(L - L_a) + E = 0 \tag{5}$$

P_i is price of price of food produced, $P_i C_i$ is a marketed surplus of produced good, P_n is the price of non-food goods, P_m is the price of a market-purchased good, W is the

wage rate, L_a is total family labor supply for on-farm use, and E is non-farm income which adjusts to ensure that Eq. (5) equals zero. Substituting the right-hand side (RHS) of Eq. (4) into 5 yields:

$$P_i(Q_i - C_i) - P_n C_n - P_m C_m - W(L - T + L_c + L_l) + E = 0 \tag{6}$$

Expanding and rearranging Eq. (6) produces an explicit household income and expenditure:

$$P_i Q_i + WT + E - WL - WL_c = P_i C_i + P_n C_n + P_m C_m + WL_l \tag{7}$$

The left-hand side of Eq. (7) represents household's full income, which comprises of the value of farm produce $P_i Q$, the value of time endowment WT , non-farm income E , the value of labor used for farming including the hired labor WL , and value of labor spent for searching and collecting scarce resources WL_c . Similarly, the right-hand side of Eq. (7) is the household expenditure on food and leisure. The expenditure side includes purchases of its own produce food consumed $P_i C_i$, value of non-food expenditure $P_n C_n$, and value of market purchase food consumed $P_m C_m$ and purchase of leisure WL_l . The optimization of Eq. (1) yields income and expenditure equation within the separability assumption. At an interior solution, the household selects L_c , L_l , L , C_i , and C_m to maximize Eq. (1) subject to Eqs. (7 and 3), which can be best visualized as:

$$\begin{aligned} \mathcal{L} = & U(C(E_f, Q_i, E, L_c; \phi), C_n, C_m, L_l; \Gamma) \\ & + \lambda(P_i Q_i + WT + E - WL - WL_c) - (P_i C_i + P_n C_n + P_m C_m + WL_l) \\ & + \gamma[F(Q_i, L, L_c, K, A)] \end{aligned} \tag{8}$$

Based on Strauss J (1983), it is possible via optimization of Eq. (8) yield production and consumption equations separately as discussed below. The first-order conditions are as follows:

$$\frac{d\mathcal{L}}{dL_c} = \frac{dU}{dC} \frac{dC}{dL_c} - \lambda W + \gamma \frac{dF}{dL_c} = 0 \tag{8.1}$$

$$\frac{d\mathcal{L}}{dL_l} = \frac{dU}{dL_l} - \lambda W = 0 \tag{8.2}$$

$$\frac{d\mathcal{L}}{dL} = \gamma \frac{dF}{dL} - \lambda W = 0 \tag{8.3}$$

$$\frac{d\mathcal{L}}{dC_i} = \frac{\partial U}{\partial C_i} - \lambda P_i = 0 \tag{8.4}$$

$$\frac{d\mathcal{L}}{dC_m} = \frac{\partial U}{\partial C_m} - \lambda P_m = 0 \tag{8.5}$$

$$\frac{d\mathcal{L}}{dC_n} = \frac{\partial U}{\partial C_n} - \lambda P_n = 0 \tag{8.6}$$

Maximizing the first-order condition of the LHS of Eq. (8) with respect to labor (L^*) and output produced (Q^*), the demand for inputs and output is derived in terms of all prices, the wage rate, time for searching and collecting scarce resource, fixed land, and capital as:

$$L^* = l^*(P_i, P_m, P_n, W, L_c, K, A) \tag{9.1}$$

$$Q^* = Q^*(P_i, P_m, P_n, W, L_c, K, A) \tag{9.2}$$

Substituting optimal labor, L^* and optimum output Q^* into LHS of Eq. (7) produces optimum income/full income Y^* under the assumption of maximized profit π^* as:

$$Y^* = P_i Q^* + WT + E - WL^* - WL_c \tag{10.1}$$

$$Y^* = WT + \pi^*(P_i, P_m, P_n, W, L_c, K, A) + E \tag{10.2}$$

where $\pi^*(P_i, P_m, P_n, W, L_c, K, A)$ represents $P_i Q^* - WL^* - WL_c$

The first-order conditions of the RHS of Eq. (7) gives consumption demand function in terms of prices, the wage rate, and income and household's preferences represented by household demographic characteristics Γ . This relationship can be specified as:

$$C_d = c(P_i, P_m, P_n, W, L_c, Y^*(P_i, P_m, P_n, W, L_c, K, A, E); \Gamma) \tag{11}$$

The above equation states that household food consumption C_d is mainly influenced by both food and non-food prices, wages, resource scarcity, and household income. Referring that household demand for food as a measure of household food security (FS), then C_d is a reduced form of the utility function in Eq. (1), which allows the evaluation of the effects of demographic and economic variables. Food security is approximated by food consumption expenditure³ in this case.

The effect of scarce resource on agricultural production is investigated through the production sector and its direct impact on household's utility is explored through consumption sector. Thus, the total effect which is the sum of the two effects can be further explained using Eq. (11). Since time spent for searching grazing or water and collecting straw is one explanatory variable of agricultural output function, the total effect of this variable on per capita food expenditure is:

$$\frac{dC_d}{dL_c} = \frac{dC_d}{dY} \frac{dY}{dL_c} + \frac{dC_d}{dL_c} \tag{12}$$

Then, the total effect is simply calculated by taking the slope coefficient of income in the consumption regression multiplied by the coefficient of time allocation in the production estimation, plus the coefficient of time allocation in the consumption regression.

Materials and methods

Study area and dataset

Ethiopia is a federal country divided into 9 regions and 2 administrative cities. Each region is subdivided into zones and zones into woredas. Woredas, in turn, are divided into peasant Associations (PA) or Tabias, an administrative unit consisting of a number of smallest villages and individual households. The study consisted of 21 Tabias stratified by agroecology and socio-economic indicators to get variations in population density and market access during the initial baseline. The main criteria used for stratification and sampling include ecology excluding lowlands (< 1500 m.a.s.l.), geographical zone (Eastern, Southern, Southeastern, Central, and Western) to reflect variations in rainfall and development pathways, distance to market based on far (> 10 km) versus near (< 10 km), population density, and irrigation access.

Initially, to reflect systematic variation in agro-climatic conditions, agricultural potential, population density, and market access conditions, four communities were selected from each of the four zones and three communities that represent irrigation projects. Likewise, one with low population density and one with high population density were strategically selected from each zone among communities to reflect far distance market (Hagos 2003). The initial data collection was carried out for a random sample of 400 households in 16 villages from the specified four zones of the region (Hagos 2003). This study is conducted in Tigray region, the northern part of Ethiopia by randomly selecting 632 sample households from 21 villages. This is because the original data set was collected by Ph.D. students of the Norwegian University of Life Science, who originally came from Tigray region. An extra 5 villages and 232 households were included to the original sample size for the simple reason that this study was conducted mainly to evaluate the impact of land certification, and thus, more control groups were required.

The data includes a panel of five rounds conducted in 1997/1998, 2000/2001, 2002/2003, 2005/2006, and 2014/2015 where the author is involved only in collecting the data for the last round. The available panel dataset provides comprehensive household and plot level data on household characteristics, agriculture and livestock information, food consumption, rental market participation, land certificate perception, and community-level data on GPS information including rainfall, total cultivated, irrigated and grazing area, wages, and conservation activities under safety net activities. This study used cross-sectional data from NMBU-MU⁴ Tigray Rural Household Survey dataset collected in 2015. The primary data used in this paper is adapted from the last, 2014/2015, household survey. The need for information regarding livestock activity reduced further our sample size from 632 to 518 only livestock owner-farmers for this study.

Econometric model specification

In order to estimate welfare, we are forced to approximate by per capita food consumption expenditure (PCFE) due to limited data.⁵ Assuming that the demand equation from the utility maximization of the recursive household model has a functional form of log-linear, its capability of estimating respective elasticities as its coefficient and modeling nonlinear effects makes it applicable and preferable (Oum 1989). Oum added that the log-linear demand function resembles the demand function obtainable from a Cobb-Douglas utility function with the drawback of invariant estimated elasticities across all data points. The aggregate demand equation per household is estimated for PCFE rather than estimating single demand equations for each product consumed or for each individual member of the household. Following Adewuyi et al. (2009); Badalona and Isitor (2014), the implicit form of the OLS is given as:

$$\ln C_d = \delta + \delta Y_i + \sum_{k=1}^K \beta_k \ln X_i + \nu \quad (13)$$

where $\ln C_d$ is households' PCFE; Y_i is rural farm and off-farm income; X_i for $k = 1 \dots K$, includes consumption side variables and household characteristics; X_1 is aggregate monetary value of crop production; X_2 is herd size in tropical livestock unit; X_3 is family size; X_4 is gender of the household head with male being equal to 1; X_5 refers to the access of information via radio, TV, and mobile in binary form; X_6 reflects the agro

ecological location of each household measured by GPS but classified as highland if it is 2500 m.a.s.l. and lowland if it is below that; X_7 represents market distance in minutes; X_8 and X_9 correspond to the dummy exposure of animal shock in 2013 and cumulative number of shocks from 2012 to 2014; $X_{10} = 1$ if the household is reported to be orthodox; $X_{11} = 1$ if the household gets assistance from relatives and friends while X_{12} and X_{13} capture the age of house head in years and total farm income composed of farm income, off-farm income, business transfer, and safety net income. The resource scarcity is captured by the walking distance to the water source in minutes/day/trip (X_{14}), walking distance to the grazing source in minutes/day/trip (X_{15}) and walking distance to the crop residue site in minutes/trip (X_{16}) per year.

ν is an error term. Since farm and off-farm income is not randomly distributed among rural households, this variable is likely to be endogenous (Hoddinott et al. 2008), which could be caused by omitted variables, measurement error, simultaneity, or household unobservable. First, a reverse causality problem might exist, because PCFE at the household level might also influence labor productivity and thus farm productivity. Second, farm and off-farm income might be influenced by household unobservable, which can lead to correlation with the error term. In the presence of endogeneity, the use of the OLS estimator biases the effect of income (Wooldridge 2009).

In order to avoid an endogeneity bias, I adopted a two-stage least square (2SLS) approach which is the most common instrumental variable estimator (Angrist and Evans 1998) where rural farm income is instrumented by shock exposure and average rainfall of 2003–2014. This is similar to approaches that have been used by Sarris et al. (2006), Hidalgo et al. (2010), and (Abdulai and Huffman 2014) in different contexts. A shock caused by crop theft, illness, and death of a household member is expected to affect income and output negatively, thereby reducing food expenditure (Abdulai and Huffman 2014; Dercon et al. 2005).

The explanation is that farm income is to decrease with increasing any shock on crop or animal farming caused by a theft or illness of the household. Then, its effect on consumption reaches through its effect on farm income. Our justification for using rainfall is that average shortfall of rainfall influence rural farm income without directly influencing the consumption expenditure in the village. Increasing rainfall is expected to increase farm income directly but consumption indirectly through its effect on income (Hidalgo et al. 2010). With this procedure, the structural equation is specified as

$$\ln C_d = \delta + \delta^{iv} \hat{Y} + \sum_{k=1}^K \beta_k \ln X_i + \varepsilon \quad (14)$$

where $\ln C_d$ is PCFE, \hat{Y} is predicted values of the endogenous rural income variable, and ε is an error term, β is parameter coefficient of the vectors of an exogenous variable, X . To obtain income (Y), the first stage regression equation is estimated by OLS based on the following specifications:

$$\ln Y = \alpha + Z' \gamma + X' \beta + \varepsilon \quad (15)$$

where, $\ln Y$ is the total rural income of the household, γ is the parameter coefficients of the vector of the instrumental variables, Z which is assumed to correlate with income (Y) but not with the error term, ε in the structural Eq. (14). The estimated PCFE of the household, in (14) is now assumed to be unbiased. In order to estimate the effects of

water and feed scarcity across the entire distribution of the dependent variables, PCFE, and to document the heterogeneity in the way food consumption respond to these scarcity variations, an alternative quantile regression was used following Koenker and Bassett (1978) estimation approach.

Results and discussions

Descriptive statistics

Table 1 presents the summary of basic variables of 518 farm households drawn from a total of 632 sample farmers. On the welfare side, the dependent variable is per capita food consumption expenditure (PCFE).⁶ For each household, expenditure profile on the following seven food groups was recorded: (1) staple foods including cereals and pulses; (2) meat, egg, and fish; (3) dairy products; (4) fruits and vegetables; (5) fats and oils; (6) sugar and honey; and (7) miscellaneous such as tea and coffee. Likewise, the

Table 1 Descriptive and summary statistics

Variables	N = 518		
	Description	Mean	SD
Dependent variables			
FE	Monetary value of food expenditure in ETB ^a	13,571.4	19,717.4
PCFE	Monetary value of per capita food expenditure in ETB	2490	3722
Output	Monetary value of crop production ^b in ETB	41,645	87,517
FI	Food Security Index ^c	0.4826	0.5001
Independent variables			
ShadowPW	Shadow price of water	147.6	204.9
ShadowPG	Shadow price of grazing	205.0	282.0
ShadowPF	Shadow price of crop residue	12.52	18.96
WaterD	Distance to animal water source in walking minute	74.85	65.54
GrazingD	Time spent looking for grazing land in walking minute	91.12	83.44
FeedD	Time to transport crop residue in walking minute	576.55	557.87
Income	Monetary value of total income ^d	49,521	92,642
Family size	Household family size	5.873	2.413
Age	Household head age	56.83	15.20
Gender	1 = male	0.743	0.437
Education	1 = literate	0.326	0.469
TLU	Herd size in TLU	3.919	3.199
MarketD	Market distance in minute	82.30	54.79
Shocks (2012–2014)	Number of shocks due to theft, flood, death	0.577	0.826
Information	1 = access to TV, radio, and mobile	0.417	0.494
Location	1 = highland (> 2500 m.a.s.l.)	0.0637	0.244
Network	1 = support from relatives and friends	0.610	0.488
Religion	1 = orthodox and 0 Muslim	0.824	0.381
Ashock13	1 = face animal shock in 2013	0.0425	0.202

^aETB refers to Ethiopian currency where 1 USD≈23 ETB during the study year (2015)

^bIt includes crop, fruit, and vegetable production

^cA household is considered food secure if it attains at least two thirds of the average PCFE of all households and considered food insecure if it falls below that value

^dIt includes income from Agriculture, off-farm, business transfer, and safety net

dependent variable on the production side is an aggregate monetary value of all crops produced during the survey production season.

An average household has produced an average agricultural output of worth 41,645 ETB and the average total income including sales from agricultural output worth 49,426 ETB. Households, on average, spend approximately about 13,571 ETB for food with average PCFE of 2490 ETB in the year. We also construct the food security-dependent variable by classifying households into food secure and food insecure using food security index calculated by dividing the individual PCFE to two-third average PCFE of all households.⁷ Accordingly, a household is considered food secure if it attains at least two thirds of the average per capita food expenditure of all households and considered food insecure if it falls below that value.

The results in Table 1 showed that 48% of the households were food secure while 52% were food insecure given the two thirds of the average of all households is 1660 ETB. Feleke et al. (2005) documented about 40% incidence of food insecurity in Ethiopia. Regarding the scarcity indicator, we know that grazing land and water resources are challenging to value because they are not traded and have no market price. Their prices are a shadow price (Magnan et al. 2012) since shadow prices are assumed to reflect better the economic scarcity of environmental goods to a household (Cooke 1998). For this reason, as a proxy indicator for scarcity, first, we use walking distance for a single trip to measure grazing, water, and crop residue using similar approach used by Palmer and MacGregor (2009). On average, the households spend 1.25 h/day to reach a water source for animal and 1.5 h/day to search for communal grazing land, maximum time reaching up to 6 h for water site and 8 h for grazing land in the data. Besides, the average time spent on collecting crop residue by the household is 9.6 h per a single trip on foot.

Second, following Baland et al. (2010), Cooke (1998), and Mekonnen et al. (2015), we measure the shadow price of searching grazing and water as well as collecting crop residue for animal as the time taken to search grazing land and water per animal or to collect crop residue per its amount collected multiplied by the village median adjusted⁸ off-farm wage. Cooke (1998) and Mekonnen et al. (2015) use shadow price of fuelwood, leaf fodder, water, and grass to measure scarcity. In this paper, we take the wage rate at the village level, and thus, there is no variation in wages for households living in the same village. In this way, we produce a household specific shadow price of searching grazing land or water and collecting straw. Table 1 reported that the average shadow price for animal watering is about 147 ETB per day which is equivalent to the average daily rural wage rate in the region. On average, the opportunity cost of searching grazing is 205 ETB per day, which is greater than the opportunity cost of water and straw. This is not surprising, as rural farmers usually spend a huge amount of time in searching grazing than watering. As expected, the shadow price of collecting a residual crop is 12 ETB per trip.

Out of the total sample, 6.4% lives in highland parts of the region. Nearly 39% of the households report that they have been severely affected by 11 different level of shocks including, drought, pests, flood, theft, illness and death, loss of job, and home damage in the last harvesting years with a mean of 0.57 shock occurrence and 4.25% of households report having been affected by animal shocks 1 year before the harvesting season. Seventy-four percent of the households are male heads with an average age of 57 years

and family size of 5.87. Since resources are very scarce, high family size may put much more pressure on consumption than it contributes to production.

Nearly 32% of the household heads have at least one or more years of education. Thus, it is hypothesized that education is negatively related to consumption value.

Around 82% of the households are Orthodox followers while 18% of the households are Muslim households in the study area. Out of the 518 households in the sample, 61% got assistance either from their relatives or friends and is expected to increase production and consumption (Di Falco et al. 2011). More than 40% of household heads site attend media via TV, radio, and mobile phone about any development intervention. Hence, it is expected that households with information are more likely to produce more and be food secure. The expected effect on production and consumption is positive (Di Falco et al. 2011). In addition, the average livestock endowment of the sample households is 4 Tropical Livestock Unit (TLU) which expected to increase food security and food consumption. Physical access to the market is measured by the amount of time required to get to the nearest local market and its mean values is reported to be 82 min. Thus, its expected effect on consumption is negative, indicating that longer distance leads to less frequency of visit and hence less likely to get market information about selling and buying prices (Feleke et al. 2005).

Econometric model results

The PCFE is analyzed using the demand functions derived from maximized utility subject to budget constraint and technology constraint of farm production, and its estimated result is presented in Tables 3 and 4 where walking distance and shadow prices are used as scarcity indicators using naive OLS and IV methods. In the IV 2SLS, total rural income is instrumented by shock exposure and average rainfall of 2003–2014. Tables 3 and 4 compare results from naive OLS and 2SLS estimates for all variables of interest, namely water, grazing land, and crop residue distance. The potential candidate instruments used in the estimation were tested to check if they could pass the necessary requirements for an instrument to be as an instrument.

Table 2 reports test results for all scenarios presented in Tables 3 and 4. The Wu-Hausman *F* test with a *P* value less than 0.05 rejected the null hypothesis that OLS estimation is consistent or income is exogenous and motivates the use of instruments. Besides, the Sargan chi-squared test fails to reject the null hypothesis that all instruments are uncorrelated with the error term in the structural model or all instruments are valid and this helps to conclude that the instruments pass the over-identification requirement for all estimates. Finally, instruments were also tested if they could pass the second most important criteria that the instrument should be correlated or relevant to

Table 2 Instrumental variables tests

Estimates	Endogeneity	Validity	Relevance
	Criteria		
	Wu-Hausman (<i>P</i> value)	Sargan (<i>P</i> value)	Stock and Yogo, <i>F</i> value
Water scarcity model	(0.0008)	(0.5562)	42.28
Gazing scarcity model	(0.0011)	(0.5236)	42.27
Straw scarcity model	(0.0013)	(0.5417)	42.56

Table 3 IV estimation of log per capita food expenditure using walking distance

Variables	(OLS)	(IV)	(OLS)	(IV)	(OLS)	(IV)
	lnPCFE	lnPCFE	lnPCFE	lnPCFE	lnPCFE	lnPCFE
Ln(output)	0.0940*** (0.0121)	0.0629*** (0.0164)	0.0909*** (0.0125)	0.0631*** (0.0164)	0.0986*** (0.0122)	0.0685*** (0.0164)
Ln(livestock)	0.0336*** (0.0129)	0.0287** (0.0134)	0.0334** (0.0131)	0.0298** (0.0134)	0.0352*** (0.0130)	0.0305** (0.0135)
Ln(Family size)	- 0.385*** (0.0529)	- 0.362*** (0.0551)	- 0.397*** (0.0535)	- 0.374*** (0.0554)	- 0.388*** (0.0534)	- 0.366*** (0.0554)
Gender(1/0)	- 0.119** (0.0588)	- 0.136** (0.0608)	- 0.0993* (0.0590)	- 0.114* (0.0607)	- 0.115* (0.0593)	- 0.133** (0.0613)
Information(1/0)	0.0591 (0.0539)	0.0409 (0.0558)	0.0454 (0.0545)	0.0288 (0.0562)	0.0487 (0.0544)	0.0299 (0.0563)
Location(1/0)	- 0.0411 (0.140)	- 0.0519 (0.144)	- 0.114 (0.140)	- 0.129 (0.143)	- 0.149 (0.141)	- 0.169 (0.145)
Ln(marketD)	0.00283 (0.0337)	0.0166 (0.0350)	0.00252 (0.0340)	0.0165 (0.0353)	0.00144 (0.0340)	0.0146 (0.0353)
Ashock13(1/0)	- 0.489** (0.191)	- 0.399** (0.199)	- 0.550*** (0.192)	- 0.463** (0.200)	- 0.540*** (0.193)	- 0.457** (0.200)
Ln(shocks)	0.212 (0.198)	0.345* (0.209)	0.307 (0.199)	0.434** (0.210)	0.267 (0.200)	0.401* (0.210)
Religion(1/0)	0.121* (0.0700)	0.146** (0.0726)	0.101 (0.0705)	0.124* (0.0727)	0.115 (0.0706)	0.140* (0.0730)
Network(1/0)	- 0.0833 (0.0554)	- 0.172*** (0.0647)	- 0.0761 (0.0559)	- 0.158** (0.0647)	- 0.0729 (0.0558)	- 0.159** (0.0649)
Age(years)	- 0.000477 (0.00174)	- 0.000749 (0.00179)	- 0.000535 (0.00175)	- 0.000786 (0.00180)	- 0.000554 (0.00175)	- 0.000829 (0.00180)
Ln(income)	0.0440*** (0.00187)	0.0565*** (0.00473)	0.0433*** (0.00189)	0.0552*** (0.00476)	0.0439*** (0.00189)	0.0562*** (0.00475)
Ln(WaterD)	- 0.122*** (0.0309)	- 0.131*** (0.0320)				
Ln(GrazingD)			- 0.100*** (0.0336)	- 0.0888** (0.0347)		
Ln(FeedD)					- 0.0642*** (0.0240)	- 0.0716*** (0.0248)
Constant	6.018*** (0.291)	5.970*** (0.300)	6.046*** (0.318)	5.898*** (0.330)	5.917*** (0.305)	5.880*** (0.319)
R-squared	0.710	0.683	0.705	0.681	0.705	0.679
First stage						
Shock		- 20.132*** (2.1697)		- 20.122*** (2.1718)		- 20.140*** (2.1686)
Rainfall		0.1655** (0.0573)		0.1612** (0.0577)		0.1657** (0.0572)
Observation	496	496	496	496	496	496

P values are for slopes: ***P < 0.01; **P < 0.05, and *P < 0.10 = significant at 1%, 5%, and 10% probability levels, respectively

Table 4 IV estimation of log per capita food expenditure using shadow prices

Variables	(OLS)	(IV)	(OLS)	(IV)	(OLS)	(IV)
	lnPCFE	lnPCFE	lnPCFE	lnPCFE	lnPCFE	lnPCFE
Ln(output)	0.0998*** (0.0122)	0.0698*** (0.0164)	0.0984*** (0.0123)	0.0677*** (0.0164)	0.0842*** (0.0131)	0.0587*** (0.0165)
Ln(livestock)	0.0381*** (0.0130)	0.0337** (0.0134)	0.0368*** (0.0130)	0.0330** (0.0135)	0.0363*** (0.0129)	0.0327** (0.0133)
Ln(Family size)	- 0.388*** (0.0536)	- 0.366*** (0.0556)	- 0.380*** (0.0538)	- 0.356*** (0.0562)	- 0.379*** (0.0533)	- 0.360*** (0.0550)
Gender(1/0)	- 0.103* (0.0593)	- 0.118* (0.0612)	- 0.0846 (0.0595)	- 0.105* (0.0618)	- 0.113* (0.0591)	- 0.125** (0.0605)
Information(1/0)	0.0544 (0.0546)	0.0363 (0.0564)	0.0401 (0.0546)	0.0250 (0.0567)	0.0412 (0.0544)	0.0267 (0.0559)
Location(1/0)	- 0.0480 (0.145)	- 0.0629 (0.149)	- 0.0567 (0.141)	- 0.0910 (0.146)	- 0.126 (0.140)	- 0.139 (0.143)
Ln(MarketD)	0.00360 (0.0341)	0.0173 (0.0354)	0.000178 (0.0342)	0.0150 (0.0357)	0.00219 (0.0339)	0.0157 (0.0350)
Ashock13(1/0)	- 0.494** (0.193)	- 0.408** (0.201)	- 0.526*** (0.192)	- 0.431** (0.201)	- 0.505*** (0.192)	- 0.426** (0.198)
Ln(Shocks)	0.220 (0.203)	0.354* (0.214)	0.241 (0.199)	0.396* (0.212)	0.247 (0.199)	0.378* (0.210)
Religion(1/0)	0.119* (0.0712)	0.143* (0.0736)	0.0948 (0.0707)	0.122* (0.0736)	0.111 (0.0704)	0.132* (0.0724)
Network(1/0)	- 0.0839 (0.0567)	- 0.170*** (0.0657)	- 0.0833 (0.0564)	- 0.167*** (0.0646)	- 0.0677 (0.0557)	- 0.147** (0.0645)
Age(years)	- 0.000615 (0.00176)	- 0.000883 (0.00181)	- 0.000760 (0.00176)	- 0.000914 (0.00182)	- 0.000504 (0.00175)	- 0.000746 (0.00179)
Ln(Income)	0.0437*** (0.00189)	0.0560*** (0.00476)	0.0426*** (0.00190)	0.0558*** (0.00482)	0.0432*** (0.00189)	0.0547*** (0.00478)
Ln(ShadowPW)	- 0.0520* (0.0295)	- 0.0528* (0.0303)				
Ln(ShadowPG)			- 0.0972*** (0.0286)	- 0.0669** (0.0312)		
Ln(ShadowPF)					- 0.0525*** (0.0172)	- 0.0441** (0.0178)
Constant	5.753*** (0.300)	5.672*** (0.309)	6.052*** (0.308)	5.785*** (0.331)	5.835*** (0.283)	5.702*** (0.294)
R-squared	0.702	0.676	0.705	0.675	0.706	0.684
First stage						
Average rainfall		0.1665*** (0.0573)		0.1459*** (0.0581)		0.1629*** (0.0574)
Shock		- 20.1556*** (2.184)		- 19.932*** (2.1718)		- 20.004*** (2.1765)
Observation	496	496	496	496	496	496

P values are for slopes: ***P < 0.01, **P < 0.05, and *P < 0.10 = significant at 1%, 5%, and 10%, probability levels, respectively

the endogenous variable income. To ensure the relevance of instruments, the Stock and Yogo (2005) F test was employed and provides higher value F statistics which is extremely higher than the rule of thumb of at least greater than 10.

The first-stage regression results of two-stage least square (2SLS) which are not reported here show that both instruments have a statistical relationship with income and carry the expected sign in all scenarios (Tables 3 and 4). Household income is often a major determinant of expenditure (Babalola and Isitor 2014). Total income of the household, which has positive coefficient significantly affected PCFE. Column (1, 3, and 5) of Table 3 shows the income effect by estimating the consumption model using OLS estimator. The coefficient of income suggests that a 1% increase in income increases PCFE by around 0.044%, whereas the 2SLS result displays that a 1 % increase in total income leads to 0.059% increase in PCFE in all estimates. Because, as the income level of the household increases, the household purchasing power increases.

It turns out that this naive ordinary estimate grossly underestimates the income effect compared to effects from the IV 2SLS estimate. This implies that estimating the model using OLS is not the correct approach and ignoring these differences would bias the income effect. The findings of Babalola and Isitor (2014), Njimanted (2006), and Thirumarpan (2013) also confirm that household income is one of the key determinants of food expenditure and food security in rural areas. We also report that farm output significantly affects household food consumption. The elasticity of PCFE with respect to the gross crop value is equal to 0.063% for IV in the water scarcity estimates. Similar effects are found in the grazing and feed estimates presented in Table 3 of columns 3 to 6. This is in line with Sarris et al. (2006) who found that agricultural productivity significantly affects PCFE in Ethiopia. The coefficient's sign and statistical significance show that livestock ownership is positively correlated with PCFE, suggesting that farmers with high herd size have a higher food consumption expenditure. Studies by Dercon et al. (2005) in Ethiopia and Sarris et al. (2006) in Tanzania found a similar result.

Another significant variable is household size, leading to 0.363% decrease in PCFE for 1% increase in the number of member of the household. This result is in line with the findings of (Bezu et al. 2014; Dercon et al. 2005) in Ethiopia and Sarris et al. (2006) in Tanzania. A household with a male head has a disadvantage of 0.136% decrement in PCFE against the findings of Dercon et al. (2005) in Ethiopia. Experiencing an animal shock at least once in the previous year lowers PCFE by 0.399%, 0.463%, and 0.457% for the three cases taking the estimated value of IV in Table 3. In line to this, Dercon (2004) found that a livestock shock negatively affects PCFE in rural Ethiopia.

The coefficient of religion is 0.146% and is statistically significant, implying that orthodox households have 0.146% PCFE higher than Muslim group referring to the IV estimate which is opposite to the result of Oldiges (2012) in India. The negative and significant sign of network shows that individuals who got social supports have 0.172% less PCFE, implying that supports from relatives or friends are not adequate enough to cover food expenditure for the recipient households. A similar result was found by Sarris et al. (2006). Other insignificant variables are proximity to market (positive), information (positive), and the age of the household head (negative) in line with the study of Matchaya and Chilonda (2012) in Malawi.

The main interest of this paper is to explore how time spent for animal feed and water searching directly affect PCFE and our result is in line with the downward spiral hypothesis (Ostrom et al. 1999). Using distance indicator in Table 3, time spent looking for water and grazing land has resulted in a negative sign and it is found to be an important factor of PCFE. A 1% increase in minutes traveled to reach water and grazing land leads to a 0.131% and 0.088% decrease in PCFE, respectively, using IV. In addition, a 1% increase in minutes traveled on foot to collect crop residue leads to 0.072% decrease in PCFE. Likewise, our results from the shadow price (Table 4) indicate that scarcity of resource have an important impact on the food demand, with the expected result that an increase in the shadow price of water, grazing, and crop residue reduces PCFE by 0.053%, 0.067%, and 0.044%, respectively. This implies that the scarcity has a negative effect on household PCFE either by affecting livestock production directly, affecting crop or off-farm income via labor reallocation or through its direct impact on time leisure consumption.

Regarding gender differential effects, Table 5 presents extra information on the gender differential effect of resource scarcity on PCFE and food security using both walking distance and shadow value indicators. For the sack of simplicity, only coefficients of variables of interests are reported. As of Table 5, the results show that a highly significant difference of PCFE and food security between male and female is a result from a resource scarcity. Taking the results from the level value, an increase in the traveling distance of the water, grazing, and transporting distance of crop residue reduces men’s PCFE by 0.149%, 0.101%, and 0.078%, respectively, but has no significant effects on female’s PCFE. Likewise, a 1% increase in minutes traveled to search water, reach grazing, and collect crop residue leads to 0.080%,

Table 5 Effect of water, grazing, and feed scarcity on log PCFE and food security

Variables	PCFE			Food security		
	All	Male	Female	All	Male	Female
A. Walking distance						
Ln(WaterD)	-0.131*** (0.0320)	-0.1492*** (0.0366)	-0.0607 (0.0665)	-0.0839*** (0.0291)	-0.0803* (0.0324)	-0.0978 (0.0762)
Ln(GrazingD)	-0.0888** (0.0347)	-0.1008** (0.0406)	-0.0339 (0.0724)	-0.1190*** (0.0335)	-0.1160*** (0.0384)	-0.1507* (0.0901)
Ln(FeedD)	-0.0716*** (0.0248)	-0.0779*** (0.0287)	-0.0394 (0.0497)	-0.0708*** (0.0234)	-0.0603** (0.0277)	-0.0822* (0.0495)
Observations	496	370	126	514	382	132
B. Shadow value						
Ln(ShadowPW)	-0.0528* (0.0303)	-0.0497 (0.0352)	-0.0550 (0.0616)	-0.0594* (0.0327)	-0.0432 (0.0373)	-0.1447* (0.0878)
Ln(ShadowPG)	-0.0669** (0.0312)	-0.0706** (0.0352)	-0.0358 (0.0705)	-0.0533* (0.0307)	-0.0513 (0.0357)	-0.0864 (0.0659)
Ln(ShadowPF)	-0.0441** (0.0178)	-0.0536*** (0.0203)	-0.0082 (0.0392)	-0.0418** (0.0185)	-0.0416* (0.0215)	-0.0330 (0.0394)
Observations	496	370	126	514	382	132

P values are for slopes: ***P < 0.01, **P < 0.05, and *P < 0.10 = significant at 1%, 5%, and 10% probability levels, respectively

0.116%, and 0.060% decrease in food security for male and 0.151% and 0.082% decrease in food security for female.

A similar pattern is revealed when analyzing differential effects of resource scarcity using shadow value scarcity indicator (Table 5). These results suggest that an increase in the shadow price of grazing and straw reduces men's PCFE by 0.071% and 0.054% but does not affect female's PCFE. Moreover, an increase in the shadow price of straw and water results in a reduction of 0.042% in food security for male and 0.145% in food security for female. These results are an indication that we find a differential effect between male and female. The negative coefficients for the scarcity indicators signal low PCFE and food security, with a male are considerably likely to have less food consumption expenditure and being food insecure more as compared to female households. This goes in line with the reality in the ground that searching and collecting these scarce resources in the study area are mainly the duty of men. Male spend more labor than female to search and collect these resources.

This result agrees with the finding of Bandyopadhyay et al. (2011) whose result revealed that scarcity of biomass negatively affected rural PCFE in Malawi. Baland et al. (2010) also showed that an increase in firewood collection time by 1 h is equivalent to an income loss of about 1% in Nepal. Bhattacharya and Innes (2006) highlighted that forest degradation spurs rural poverty in Sub-Saharan Africa. This supports the argument by Chopra et al. (2007), Cooke et al. (2008), Kumar and Hotchkiss (1988), and Tangka and Jabbar (2005), whose study conclude that feed and water scarcity reduces livestock, crop, and non-farm productivity as well as access to food, resulting in less food security and low human welfare by traveling long distance with an animal in search of feed and water in less developing countries.

Estimation with regard to the food security is presented in Table 6, the model had about 38% prediction as compared to 48% observed probability. The negative significant relationship between the shadow prices and the household food security implies that household who spend more time on searching water, grazing, and crop residue are more likely to be food insecure than their counterpart with nearer distance. The coefficients from marginal effect indicated that increasing the shadow prices of water, grazing, and crop residue reduces the probability of food security by 0.0594%, 0.0533%, and 0.0418%, respectively, supporting the arguments forwarded by Cooke et al. (2008) and Mekonnen et al. (2015). The results further show that the probability of food security increases significantly and consistently with farm output, total income, and religion in favor of (Ogundari 2017) but declines with family and herd size, supporting the results from (Feleke et al. 2005).

The hypothesis that the impact of feed and water scarcity strongly increase from the bottom (poor) to the top (rich) quantile is tested using quantile regression results displayed in Table 7. Surprisingly, time spent looking for water and grazing land has resulted in a negative sign as expected and it is found to be an important factor in per capita food expenditure. The impact of a 1% increase in distance to grazing and crop residue source brings about a 0.171% and 0.069% reduction in food expenditure only for the top category while the effect of water is 0.064% at the median value. This supports the argument by Tangka and Jabbar (2005), whose study concluded that feed scarcity reduces livestock, crop, and non-farm productivity as well as access to food, resulting in less food insecurity and low human welfare by traveling long distance with the animal in search of feed and water in less developing countries.

Table 6 Probit estimation of food security using walking distance

Variables	(ME)	(ME)	(ME)
	HHFS	HHFS	HHFS
Ln(output)	0.0280** (0.0134)	0.0315** (0.0140)	0.0153 (0.0145)
Ln(livestock)	-0.0257** (0.0120)	-0.0236* (0.0122)	-0.0269*** (0.0121)
Ln(Family size)	-0.203*** (0.0591)	-0.202*** (0.0596)	-0.206*** (0.0598)
Gender(1/0)	-0.0880 (0.0670)	-0.0596 (0.0669)	-0.0853 (0.0668)
Information(1/0)	0.0629 (0.0584)	0.0358 (0.0592)	0.0506 (0.0585)
Location(1/0)	-0.0963 (0.146)	-0.126 (0.134)	-0.169 (0.124)
Ln(marketD)	-0.0447 (0.0376)	-0.0540 (0.0379)	-0.0409 (0.0374)
Shock13(1/0)	-0.138 (0.197)	-0.155 (0.187)	-0.151 (0.189)
Ln(shocks)	0.355 (0.240)	0.404* (0.219)	0.395* (0.224)
Religion(1/0)	0.147** (0.0649)	0.135** (0.0653)	0.135** (0.0653)
Network(1/0)	-0.0703 (0.0645)	-0.0419 (0.0640)	-0.0482 (0.0636)
Age(years)	-0.000700 (0.00185)	-0.000232 (0.00185)	-0.000741 (0.00185)
Ln(income)	0.0327*** (0.00293)	0.0321*** (0.00293)	0.0324*** (0.00295)
Ln(ShadowPW)	-0.0594* (0.0327)		
Ln(ShadowPG)		-0.0533* (0.0307)	
Ln(ShadowPF)			-0.0418** (0.0185)
Observed probability	0.4824	0.4792	0.4824
Predicted probability	0.3803	0.3723	0.3792
Pseudo R^2	0.4379	0.4404	0.4405
Observation	514	514	514

P values are for slopes: *** $P < 0.01$, ** $P < 0.05$, and * $P < 0.10$ = significant at 1%, 5%, and 10% probability level, respectively

Total effect of feed and water scarcity on total welfare

This analysis finalizes its discussion by exploring the total effect of animal water and feed scarcity on total welfare. Based on the descriptive statistics in Table 1, the median household in this sample spends up to one 75 min to travel to a water source, 91 min to search for grazing land, and 577 min to transport crop residue

Table 7 Effect of water, grazing, and feed scarcity on log PCFE using quantile regression

Variables	(PCFE)	(PCFE)	(PCFE)	(PCFE)	(PCFE)
	q10	q25	q50	q75	q90
Ln(ShadowPW)	-0.00210 (0.0546)	-0.0102 (0.0291)	-0.0644** (0.0259)	-0.0341 (0.0268)	-0.0299 (0.0457)
Ln(ShadowPG)	-0.0608 (0.0408)	-0.0345 (0.0274)	-0.100*** (0.0313)	-0.0996*** (0.0289)	-0.171*** (0.0490)
Ln(ShadowPF)	-0.0413 (0.0372)	-0.00441 (0.0247)	-0.00858 (0.0181)	-0.0194 (0.0165)	-0.0691*** (0.0260)
Observations	496	496	496	496	496

P values are for slopes: ***P < 0.01 and **P < 0.05= significant at 1% and 5% probability level, respectively

yearly. The labor hours allocated for these scarce resources then reduces the total time available for crop farming activities in addition to the reduction on the households’ leisure consumption. Its effect on agricultural production is investigated via the production sector and its direct impact on household’s utility analyzed through consumption sector. The aggregate of the two shows the total welfare effect on the household’s livelihood.

Based on Eq. (12), the total effect is simply calculated by taking the slope coefficient of income ($\frac{dC_d}{dY}$) in the consumption regression multiplied by the coefficient of time allocation in the production estimation($\frac{dY}{dL_c}$), plus the coefficient of time allocation in the consumption regression ($\frac{dC_d}{dL_c}$). Based on Table 8, the total impact of time spent searching for water, feed, and collecting straw on per PCFE is -0.142%, -0.102%, and -0.092%, respectively, using distance measure. This implies that for a 1% increase in minutes traveled to a water, grazing, and straw source, PCFE decreases by 0.142%, 0.102%, and 0.092%, respectively. If the median household in this data spends about 60 min daily to look for water and feed source and have PCFE 2490 ETB, decreasing traveling minutes to a water, grazing, and straw source by 0.6 min daily will increase PCFE by 354 ETB, 254 ETB, and 229 ETB, respectively, for the median household using panel A distance value (Table 8). The estimate for the production is available upon request.

Table 8 Aggregate effect of water and feed scarcity on output, food expenditure, and food security

Estimates	Effect on output (Y)	Effect on PCFE	Total effect
Panel A using distance value	$\frac{dY}{dT}$	$\frac{dPCFE}{dT}$	$\frac{dPCFE}{dY} \cdot \frac{dY}{dT} + \frac{dPCFE}{dT}$
Water scarcity (T_w)	-0.155	-0.133	-0.142
Grazing scarcity (T_g)	-0.279	-0.086	-0.102
Straw scarcity (T_s)	-0.328	-0.0731	-0.092
Panel B using shadow price			
Water scarcity (T_w)	-0.074	-0.0529	-0.057
Grazing scarcity (T_g)	-0.094	-0.0627	-0.068
Straw scarcity (T_s)	-0.154	-0.0421	-0.051

Conclusion and suggestion

The scarcity of grazing and water for an animal has a negative effect on households' welfare and food security either by affecting livestock production directly, affecting crop or off-farm income via labor reallocation or through its direct impact on time leisure consumption. Our research questions focus on the relationship between natural resource scarcity and PCFE (welfare) and food security (Table 8). In this paper, we have explored these effects using distance and shadow price as resource scarcity indicators in Northern Ethiopia based on 518 sample farmers. To address our research first objective, we employed the IV-2SLS estimation, and the second question is addressed by estimating a probit model for food security. The descriptive result shows that about 48% of the households were food secure while 52% were food insecure given the two thirds of the average of all household's PCFE is 1660 ETB.

Our results confirmed the theoretical prediction that resource scarcity affects households' welfare and food security adversely as predicted by the downward spiral hypothesis (Ostrom et al. 1999). The results in this paper provide an interesting picture of smallholders in Ethiopia and hint at several areas that could be important for improving food security and welfare in general. As expected, it appears that time spent looking for water and feed has a significant and negative effect on PCFE and food security. In aggregate, reducing time spent looking for water by 1% leads to an increase in PCFE by 0.131% and food security by 0.0594%. Similarly, a 1% decrease in time wastage for searching grazing land increase PCFE and aggregate food security by 0.088% and 0.053%, respectively, and an increment of 0.0716% in PCFE and 0.0418% in food security is achieved by 1% reduction in crop residue transporting time per a single trip.

The total impact of time spent searching for water, feed, and collecting straw on per PCFE is -0.142% , -0.102% and -0.092% , respectively, using distance measure, and decreasing traveling minutes to a water, grazing, and straw source by 0.6 min will increase PCFE by 354 ETB, 254 ETB, and 229 ETB for the median household. Depending on results from the quantile regression, the effect of water and feed scarcity is not uniform across the food consumption distribution. A similar pattern is revealed when analyzing differential effects of resource scarcity. The results suggest that an increase in the traveling distance of water, grazing, and transporting distance of crop residue reduces men's PCFE by 0.149%, 0.101%, and 0.078% but have no significant effects on female's PCFE. Likewise, a 1% increase in minutes traveled to search the water, reach grazing, and collect crop residue leads to 0.080%, 0.116%, and 0.060% decrease in food security for male and 0.151% and 0.082% decrease in food security for female. The negative coefficients for the scarcity indicators signal low PCFE and food security, with a male are considerably likely to have less food consumption expenditure and being food insecure more as compared to female households.

Our study plays a great role in the understanding of the linkages between welfare, food security, and environmental resources such as grazing and water scarcity. In general, this study can be helpful for policymakers working to alleviate animal water and feed problems in Ethiopia to justify their actions with an empirical result. Besides, this study's result can give a good lesson for policy analysts that labor allocation for reaching water and feed source imposes a negative impact on on-farm farmers' agricultural output and food consumption and hence on food security by displacing productive labor away from productive activities. Since the effect of these scarce resources is not

uniform between poor and rich or between male- and female-headed farmers. A nearby water and feed source does not only alleviate labor constraints, but also saves energy and time that could be used for other productive farming activities.

Two areas of policy intervention can be emerged as relevant. The first involves policies and institutions that facilitate easier access to animal water tap by advocating on emergency relief grounds. The second area of policy intervention involves the introduction of more efficient animal feed management strategy that can improve cattle production and reduce land degradation. A policy that sought to increase aggregate production or that which sought to increase household food consumption would greatly impact the highest quantile more than those who are in the lowest quantile of food production distribution. Future research should focus on adopting an approach using other welfare indicators (nutrition, incidences of malnourishment, daily calorie intake, and diversity of food intake) to proxy food security than per capita food expenditure. In addition, a research that uses a lag value of crop output than current output in the per capita food expenditure estimation based on longitudinal data will be a plus for this paper.

Endnotes

¹See for a similar approach in the work of (Baland et al. 2010; Cooke 1998; and Cooke et al. 2008).

²For a detail review of related empirical studies, see Cooke et al. (2008).

³See for a similar approach in the work of (Smith and Subandoro 2007; Gaiha et al. 2014; Mignouna et al. 2015; Çağlayan and Astar 2012).

⁴NMBU-MU refers to the Norwegian University of Life Science-Mekelle University.

⁵Check the work of Asfaw et al. (2012) and Thirumarpan (2013) for similar work.

⁶Thirumarpan (2013) and Asfaw et al. (2012) used consumption expenditure to reflect the socio-economic welfare of household and is a reliable indicator of food accessibility and degree of vulnerability to food insecurity.

⁷The same approach is found in the work of Titus and Adetokunbo (2007).

⁸In order to adjust for big variation in the wage rate among villages of the region, the wage rate is adjusted using a general informal rural labor conversion factor, 0.98.

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Availability of data and materials

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Author's contributions

The single author revised and approved the final submitted manuscript.

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