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## Article

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Article

# Investigating the Dynamic Interlinkages between Exchange Rates and the NSE NIFTY Index

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**Abstract:** This study aims at examining the short-run and long-run dynamic linkages among exchange rates and stock market index in India through a structured cointegration and Granger causality tests. Daily exchange rates of USD, EUR, CNY, JPY, and GBP to INR along with the daily movement of NSE NIFTY for a period spanning 13 years from 6 September 2005 to 31 December 2018 were used for the analysis. The results reveal that there is no evidence for a stable long-run relationship between NSE NIFTY and the exchange rates under study. However, the VAR-based Granger causality test shows that USD, JPY, and CNY have short-run causal relationship with NSE NIFTY. The NSE NIFTY also seemed to have an influence on USD expressed in terms of Indian rupee. The impulse response analysis further supports the results of the Granger causality test and provides information on the time required for the NSE NIFTY index to recover from a shock caused by the fluctuation in exchange rates.

**Keywords:** NIFTY; currencies; Granger causality; impulse response



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## 1. Introduction

Fluctuation in exchange rates is one of the significant factors that affect stock prices, which subsequently influences a firm's market value. Despite the fact that this topic has been discussed worldwide, there is still a lack of unanimity in the literature regarding the relationship between exchange rates and stock market indices (Mroua and Trabelsi 2020). As financial theory describes, the interest rates and exchange rates significantly influence the value of a firm and the upward and downward movements of the exchange rate play a pivotal role in determining the stock prices. In both developed as well as developing countries, the stock market plays a crucial role as far as financial intermediation is concerned. It channels the funds in an economy from surplus units to deficit units through resource mobilization, which is critical in the process of expansion and growth of an economy. Hence, the stock market acts as a channel in mobilizing savings and helps in the efficient allocation of funds, thereby fostering economic growth (Olugbenga 2012; Alile 1984).

Various hypotheses support the causal relationship between exchange rates and stock prices, such as, goods market approach, which defines that the competitive nature of the firm is affected by the change in exchange rates. It signifies that the exchange rate fluctuations affect the earning value and cost of funds along with their stock prices because companies start to borrow in foreign currencies in order to fund their processes (Dornbusch and Fischer 1980). The export of goods becomes attractive when there is a depreciation in the local currency, and it leads to an increase in foreign demand. As a result, the revenue of

a firm appreciates along with the appreciation in the firm's value and stock prices. Similarly, when the local currency appreciates, it leads to a decrease in profits of a firm that exports goods as there is a decrease in its products' foreign demand. However, there is a change in the sensitivity of the value of the firm to the change in the exchange rate, and it depends on whether it is an importing firm or an exporting firm (Nath and Samanta 2003).

An alternative approach is provided by the portfolio balance models, which stressed the role of transactions in capital account in studying the relationship between stock prices and exchange rates. For example, capital inflows from foreign investors have been drawn by an attractive performance of a stock market, which increases the demand for domestic currency. On the other hand, the demand for the local currency depreciates when there is a fall in stock prices as the foreign investors start selling their stocks (Adebiyi et al. 2009). Moreover, with time, foreign investment can increase in local equities as a result of the profits of international broadening to foreign investors. Hence, the upward and downward movements of stock prices affect the money demand and exchange rates because the liquidity demand and wealth of investors could be a factor that affects the stock market performance (Mishra 2004).

Several theories have suggested a causal relationship between the exchange rates and stock prices. Nevertheless, evidence that provides mixed results is also widely available. For instance, out of 171 multinational firms of Japan, only 25 Percent have shown a significant disclosure of exchange rate over the stock market (He and Ng 1998). Another study showed the empirical results that there was only a negligible impact of weekly exchange rate movement on the value of industry indices in many countries (Griffin and Stulz 2001). Moreover, some theories failed to identify a significant relationship between the movement of dollar prices and stock returns of US firms, for instance, Jorion (1990); Bartov and Bodnar (1994), etc.

Considering the interrelationship between the stock market indices and exchange rates, it is imperative that more studies need to be conducted in this regard. However, there is still a dearth of comprehensive studies that investigate this dynamic interrelation in the Indian context. The main aim of this study is to examine the linkages between the exchange rates and NSE NIFTY index using the daily data of NSE NIFTY closing values and exchange rates spanning 13 years (2005–2018). The objective is to identify whether variations in exchange rates influence the NSE NIFTY closing values, which can be a significant cue for the investors while making investment decisions. The NIFTY 50 is an index that indicates the weighted average of fifty Indian companies that are listed on the National Stock Exchange of India (NSE), which is the largest stock market in the country. This study analyses the long-run relationship between the exchange rate of US Dollar (USD), Euro (EUR), Chinese Yuan (CNY), Japanese Yen (JPY), British Pound (GBP) to Indian Rupee (INR) with NSE NIFTY index using a Johansen cointegration test. The short-run relationship between the above-mentioned variables is examined using a VAR Granger Causality test and impulse response function.

## 2. Literature Review

In a study conducted by Ibrahim (2000), the interactions between stock prices and exchange rates in Malaysia were examined using a Cointegration and Granger causality test. The study found no long-run relationship between stock prices and exchange rates, but evidence for cointegration was observed when M2 money supply and reserves were included in the analysis. The study construed that the short-run exchange rates did impact the stock market price. Likewise, other studies also supported that there is no significant relationship between the exchange rate and stock market indices in the long-run (Frank and Young 1972; Bhattacharya and Mukherjee 2002).

Phylaktis and Ravazzolo (2005) analyzed the short-run and long-run effects between stock prices and exchange rates and the channels through which exogenous shocks impact the markets of some Pacific Basin countries for a period between 1980 and 1998. A structured cointegration and multivariate Granger causality test were used for the analysis and

the results showed a positive long-run and short-run causality between stock prices and exchange rates. Another study by [Gulati and Kakhani \(2012\)](#) to find a causal relationship between INR/Dollar exchange rates and stock market indices (SENSEX and NIFTY) for a period of 2004–2012 used Granger causality and correlation analysis. Their results showed that there is a weak positive correlation between stock price indices and exchange rates for the aforementioned period.

The nature of the causal relationship between exchange rates and stock market performances of Switzerland and Poland from 2001 to 2008 was examined using both linear and non-linear causality tests. The results showed that the performance of the stock market is a causal factor for changes in exchange rates (Portfolio approach) for Switzerland. For Poland, both portfolio and traditional approach where changes in exchange rates caused fluctuations in stock market prices had an impact on the exchange rates ([Gurgul and Lach 2012](#)).

A study conducted in Nigeria for the period between 2001 and 2011 shows no significant relationship between exchange rate and stock market prices. The result of the Granger causality test indicates the absence of causality between the All Share Index (ASI) and exchange rates, which show the independence of exchange rates and stock market index in Nigeria ([Zubair 2013](#)). Similarly, a study conducted by [Lee and Zhao \(2014\)](#) explained the short-run and long-run causal relationship between stock market prices and exchange rates in China for a period between 2002 and 2012 using cointegration tests, vector error correction (VEC) estimates, block exogeneity Wald tests, impulse responses, variance decomposition techniques, and structural break tests. The results showed a long-run causality from exchange rates to stock prices in Chinese stock markets and a short-run causality from Japanese yen and Korean won exchange rates to stock prices in Shanghai Stock Exchange strongly prevails, whereas in the Shenzhen Stock Exchange the causality is weak. The impact of the global financial crisis from 2007 to 2009 on the Chinese stock markets was found to be insignificant.

[Chkili and Nguyen \(2014\)](#) studied the upward and downward movement of exchange rates and stock market returns in a regime-switching environment for BRICS countries for 16 years (1997–2013) using Markov autoregressive model and VAR model. The results showed that the returns from the stock market were higher than exchange returns in all BRICS countries. Among BRICS countries, South Africa is less volatile, and Russia is more volatile and changes in exchange rate do not affect the stock market return of BRICS countries. Inversely, the impact from stock market returns to exchange rates is significant for all BRICS countries. Likewise, a study conducted for Karachi stock exchange (KSE 100 Index, Pakistan) using the Johansen cointegration test and Granger causality test showed that there was no long-run relationship between exchange rates and the stock market prices ([Ihsan et al. 2015](#)). [Poornima and Ganeshwari \(2016\)](#), in their study about stock exchange in India, found that there exists a unidirectional relationship between exchange rates and the NIFTY Index.

Another study about the Indian stock market and foreign exchange rates of USD, Euro, Yen, and Pound sterling against the Indian rupee (INR) was conducted for a period of five years, 2011–2015. The study used the Granger causality test to find the cause and effect relationship between the stock market index and the exchange rates. The study explains that the cause and effect between stock market indices (Sensex and NIFTY) shows an inverse relationship between returns from foreign exchange rates USD, GBP, EURO, and YEN and the relationship between currency returns to other currency returns against INR shows a positive relationship. This explains that the cause and effect relationship is unidirectional ([Kumarasamy and Chellasamy 2017](#)).

From the literature, it could be seen that most of the studies in the related area used cointegration and Granger causality to test the long-run and short-run relationship between exchange rates and stock market indices. This study using similar methods hypothesizes that:

**Hypothesis 1.** *There is a cointegrating relationship between the exchange rates and the NSE NIFTY index.*

**Hypothesis 2.** *USD Granger causes NSE NIFTY index.*

**Hypothesis 3.** *EUR Granger causes NSE NIFTY index.*

**Hypothesis 4.** *CNY Granger causes NSE NIFTY index.*

**Hypothesis 5.** *JPY Granger causes NSE NIFTY index.*

**Hypothesis 6.** *GBP Granger causes NSE NIFTY index.*

### 3. Methodology

#### 3.1. Data Description

The data used for the study were collected from the Bloomberg Terminal and National Stock Exchange (NSE) website. Daily exchange rates of USD, EUR, CNY, JPY, and GBP to INR along with the daily movement of NSE NIFTY 50 for a period spanning 13 years from 6 September 2005 to 31 December 2018 comprising 3286 observations were used for the analysis.

#### 3.2. Cointegration

Regression analysis using non-stationary time series often makes the results spurious. Taking the first difference may not make the variables suitable for regression in most cases as information regarding co-movement may be lost in the process and ultimately result in poor forecasting (Sims 1980). In order to solve this issue, Engle and Granger (1987) suggested testing for cointegration among non-stationary time series. If the variables are cointegrated of the same order, then an “error correction” is said to exist, which explains the short-run dynamics between the cointegrated variables. In such cases, the appropriate method is to estimate the Vector Error Correction Model (VECM), which is the same as the Vector Auto Regression (VAR) in the first difference with a vector of cointegrating residuals. If the variables are not cointegrated, then a VAR model can be fitted.

#### 3.3. Granger Causality Test

A time series  $Y_t$  Granger causes another time series  $X_t$ , if the present values of  $X$  can be predicted using the past values of  $Y$  and other relevant information, which includes the past values of  $X$  as well (Granger 1969). An F statistic is commonly used to interpret the results of causality. In the case of using more than two variables, such as an additional variable  $Z$ , which also mutually causes  $X$  or  $Y$ , a multivariate framework could be used. The Granger causality test used in this study is carried out in a VAR framework, which provides a flexible environment to conduct the analysis. The methods followed in this study are based on the procedure given by Granger (1969) and Engle and Granger (1987). Nevertheless, the specific technique to be used mainly depended upon the presence of cointegration among the variables.

#### 3.4. Impulse Response Function

An impulse response function in very simple terms could be explained as a technique used to examine the path of a variable of interest in a VAR model after encountering a shock. The transmission of a shock could be tracked with the help of an impulse response function which facilitates the proper assessment of the behavior of a variable (Alloza 2017). Two methods are commonly used to depict the dynamic behavior of a VAR model and to identify the sources of variability: Impulse Response Function (IRF) and Variance Decomposition. The variance decomposition method suffers from limitations that may affect the characteristics of impulses. The generalized impulse response functions propounded by Pesaran and Shin (1998) offer better results as compared to the former. This study uses the

generalized impulse response function to analyze the dynamic behavior of the variables.

#### 4. Results and Discussion

##### 4.1. Descriptive Statistics

The descriptive statistics, along with the normality test for all the variables under study, are given in Table 1.

**Table 1.** Descriptive Statistics and Normality Test.

Descriptive Statistics	NSE	USD	EUR	CNY	JPY	GBP
Mean	6256.61	54.2226	68.5863	8.12960	0.53318	85.0561
Std. Dev	2298.33	9.79310	8.73385	1.84054	0.09875	9.36010
Skewness	0.46153	0.15552	0.07112	−0.14761	−0.51928	0.23232
Kurtosis	2.28362	1.54340	2.09774	1.45031	2.26508	2.19437
Wilk-Shapiro	0.91757	0.93570	0.97576	0.93865	0.95046	0.95835
<i>p</i> -Value	0.00000	0.00000	0.00000	0.00000	0.00000	0.00000

Note: SW test null—Normal Distribution.

The descriptive statistics results include the mean values showing the average daily exchange rate variations, standard deviation indicating the extent of volatility from the mean value, skewness and kurtosis for normality tests are given in the table above. The Shapiro–Wilk normality test was also used to test the normality of the time series distributions. As can be seen, the kurtosis values are higher than the specified threshold, which implies that the time series data do not follow normal distribution. This is further confirmed using a Shapiro–Wilk Test, which rejects the null hypothesis that the data follow a normal distribution.

##### 4.2. Stationarity Test

In order to test for cointegration, it is important to check the order of integration. The results of both Augmented Dickey Fuller (ADF) test and Kwiatkowski–Phillips–Schmidt–Shin (KPSS) test show that the variables are not stationary at level but are stationary after first differencing. Table 2 shows the results of the stationarity test.

**Table 2.** Tests for Stationarity.

Variable	Levels		First Difference ( $\Delta \ln$ )		Inference
	Dickey Fuller	KPSS	Dickey Fuller	KPSS	
NSE	−2.6286	28.66 **	−14.14 **	0.063	I(1)
USD	−2.7225	30.35 **	−13.78 **	0.085	I(1)
EUR	−2.6925	23.73 **	−14.78 **	0.041	I(1)
CNY	−1.8834	30.81 **	−13.64 **	0.156	I(1)
JPY	−2.0957	20.89 **	−13.96 **	0.088	I(1)
GBP	−2.0798	13.56 **	−14.59 **	0.065	I(1)

Note: Significant at 0.05, \*\* Significant at 0.01.

Line plots for the exchange rates and NSE NIFTY index for the period under study are shown in Figure 1.

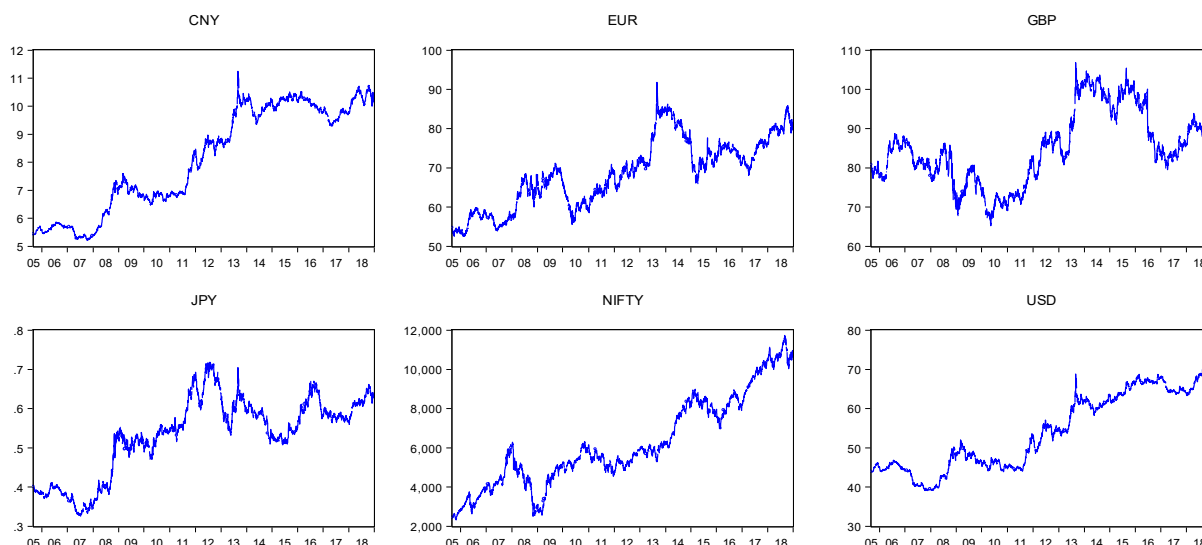


Figure 1. Line plots of the variables over the study period.

4.3. Test for Cointegration

Now that the variables follow the same level of integration, the existence of cointegration among them can be tested. If two variables  $x$  and  $y$  are  $I(1)$  processes and their linear combination  $et = \beta_0 + \beta_1 X + \beta_2 Y$ , is stationary then the variables are said to be cointegrated of the order  $(1,1)$ . The stochastic trends of these variables linked, such that they cannot drift apart in the long-run, pointing to an equilibrium relationship between them.

A Johansen multivariate cointegration test (Johansen 1988) was carried out to identify a cointegrating relationship between the variables. The test statistic, critical values, and rank of cointegration are shown in Table 3.

Table 3. Johansen Cointegration—Trace Statistic with linear trend.

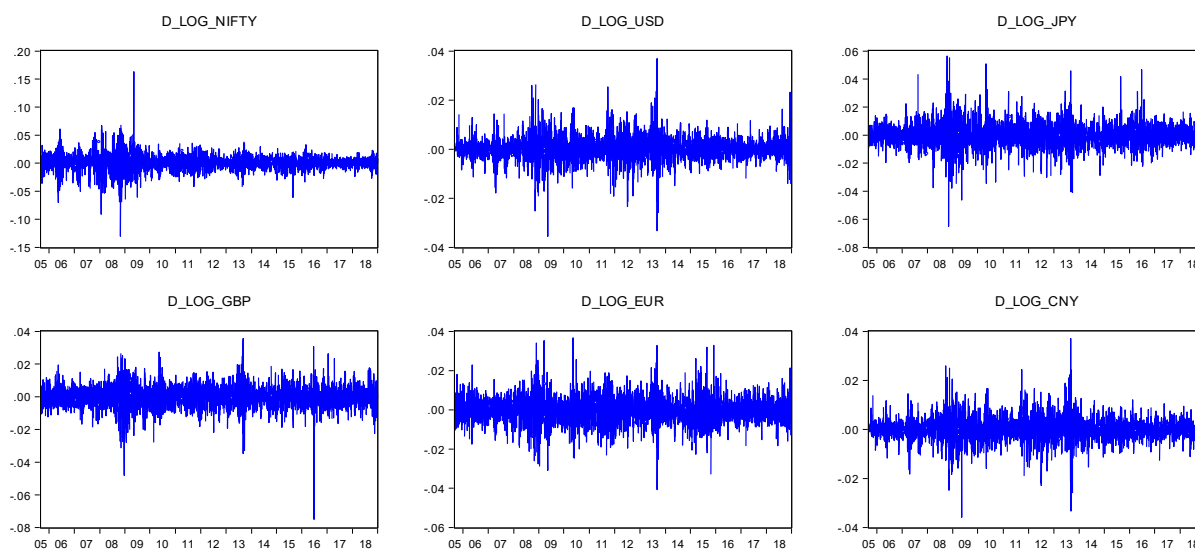
Rank	Test Statistic	Critical Values		
		10 pct	5 pct	1 pct
$r \leq 5$	1.17	6.50	8.18	11.65
$r \leq 4$	5.50	15.66	17.95	23.52
$r \leq 3$	13.27	28.71	31.52	37.22
$r \leq 2$	22.52	45.23	48.28	55.43
$r \leq 1$	35.34	66.49	70.60	78.87
$r \leq 0$	64.41	85.18	90.39	104.20

The test results reject the hypothesis that the variables are cointegrated as the test statistic value for  $r \leq 0$  cannot be rejected at 10 pct, 5 pct, and 1 pct. The test statistic values are not higher than the critical values for other cointegrating ranks as well. The results imply that there exists no long-run relationship between the exchange rates and the NIFTY index. Hypothesis 1 is rejected as no evidence could be found for a cointegrating relationship between the variables. The finding that USD to INR does not have a long-run relationship with NSE NIFTY index is corroborated by Bhuvaneshwari and Ramya (2017) using Johansen Cointegration test.

4.4. Causality Test

After finding no evidence for long-run relationship between the variables, the short-run causation effect was investigated using the Granger causality test. Stationarity of the variables is a prerequisite for the Granger causality test. Hence the log difference of the

variables was used to ensure that the variables follow constant statistical properties over the period of analysis. A unidirectional causality test using Granger causality test in VAR environment was carried out to check whether the changes in the exchange rate can predict the NIFTY index. The graphical representation of the variables in their first log difference is shown in Figure 2.



**Figure 2.** Variables after first log difference.

This study's scope is confined to the assessment of the impact of various exchange rates on the movement of NSE NIFTY index. Nevertheless, NSE NIFTY's influence on the exchange rates that are expressed in terms of Indian rupee was also estimated to identify causal effects. The lag order was set based on AIC criteria. The 13 years were split into three specific time durations (September 2005–September 2009, October 2009–October 2013 and November 2013–December 2018) to assess the short-term influence of movements in exchange rates on NSE NIFTY. Further to this, Granger causality test was also run for the overall period.

#### 4.4.1. Split Up Periods

The split wise Granger causality test results show that the various currency pairs drive the nifty index in different periods. During the first split (2005–2009), the Euro and JPY were seen causing Nifty 50 movements. Nifty 50 was, in turn, Granger causing USD and CNY. In the Second split (2009–2013), Euro was seen as Granger causing Nifty and Nifty was found Granger causing Euro, JPY and GBP. These movements may be primarily contributed to the foreign trade figures. European Union, China and the USA were among the top 10 trading partners to India during those periods ([Trade Statistics by Country | WITS n.d.](#)).

The Final Split (2013–2018) further shows that USD, CNY and JPY are causing Nifty and Nifty causing USD and CNY. Along with the foreign trade figures (with China, USA, and Japan) the FII flows are also the reason for the bi-directional relationship between USD and Nifty. Overall, the research found that the short-term relationships between the various currency pairs and nifty index is primarily driven by the foreign trade figures and FII flows. Tables 4–6 show the granger causality test results for the split periods; September 2005 to September 2009, October 2009 to October 2013 and November 2013 to December 2018 respectively.



**Table 4.** Granger causality in VAR Framework for the period September 2005–September 2009.

Independent Variables	Dependent Variable	F Statistic	p-Value	Independent Variable	Dependent Variable	F Statistic	p-Value
$\Delta \ln \text{USD}$	$\Delta \ln \text{NSE}$	1.766	0.184	$\Delta \ln \text{NSE}$	$\Delta \ln \text{USD}$	8.457	0.003
$\Delta \ln \text{EUR}$		3.060	0.015		$\Delta \ln \text{EUR}$	0.276	0.893
$\Delta \ln \text{CNY}$		1.546	0.214		$\Delta \ln \text{CNY}$	7.248	0.007
$\Delta \ln \text{JPY}$		12.50	0.000		$\Delta \ln \text{JPY}$	0.223	0.636
$\Delta \ln \text{GBP}$		0.008	0.926		$\Delta \ln \text{GBP}$	0.180	0.670

Note: Granger Causality Null—No Granger Causality.

**Table 5.** Granger causality in VAR Framework for the period October 2009–October 2013.

Independent Variables	Dependent Variable	F Statistic	p-Value	Independent Variable	Dependent Variable	F Statistic	p-Value
$\Delta \ln \text{USD}$	$\Delta \ln \text{NSE}$	1.225	0.293	$\Delta \ln \text{NSE}$	$\Delta \ln \text{USD}$	1.667	0.189
$\Delta \ln \text{EUR}$		5.846	0.015		$\Delta \ln \text{EUR}$	8.446	0.003
$\Delta \ln \text{CNY}$		1.268	0.210		$\Delta \ln \text{CNY}$	1.657	0.190
$\Delta \ln \text{JPY}$		0.281	0.754		$\Delta \ln \text{JPY}$	4.433	0.011
$\Delta \ln \text{GBP}$		0.594	0.552		$\Delta \ln \text{GBP}$	7.145	0.000

Note: Granger Causality Null—No Granger Causality.

**Table 6.** Granger causality in VAR Framework for the period November 2013–December 2018.

Independent Variables	Dependent Variable	F Statistic	p-Value	Independent Variable	Dependent Variable	F Statistic	p-Value
$\Delta \ln \text{USD}$	$\Delta \ln \text{NSE}$	4.636	0.000	$\Delta \ln \text{NSE}$	$\Delta \ln \text{USD}$	3.706	0.002
$\Delta \ln \text{EUR}$		3.336	0.067		$\Delta \ln \text{EUR}$	0.000	0.982
$\Delta \ln \text{CNY}$		2.507	0.007		$\Delta \ln \text{CNY}$	2.225	0.018
$\Delta \ln \text{JPY}$		5.464	0.019		$\Delta \ln \text{JPY}$	0.277	0.598
$\Delta \ln \text{GBP}$		0.932	0.334		$\Delta \ln \text{GBP}$	0.2565	0.612

Note: Granger Causality Null—No Granger Causality.

#### 4.4.2. Overall Period (2005–2018)

The results of the Granger causality test (from 2005 to 2018) show that the changes in the USD, CNY, and JPY Granger cause NSE NIFTY implying that fluctuations in the exchange rate of USD, CNY, or JPY will cause volatility in NSE in the short-run. However, other exchange rates did not seem to have a causal relationship with the NSE NIFTY index. The USD/INR exchange rates and nifty movements are interconnected, which re-establishes the fact that Indian markets are still tracking the US economy and substantial dollar-denominated funding happens in the Nifty index. The dollar-denominated transactions further point out that the fund flows which gained momentum after quantitative easing (after 2009) have not stopped yet. This may mean that the Indian stock markets are still giving better spread in comparison to US domestic investments. This is evident from the latest charts of the United States and Indian 10 year Government securities. As of the 24 August 2020, Indian 10 year was trading at 6.22 levels and the US 10 year was at 0.64 levels ([Trading Economics n.d.](#)).

Another reason for the causal relationship between the above-mentioned exchange rates and the NIFTY index is the trade relations of India with these countries. The US became India's largest trade partner with a trading volume of \$87.95 billion in 2018–2019 by surpassing China with which the bilateral trade was worth \$87.07 billion ([The Economic Times 2019](#)). The causality between the trade-weighted dollar index ( $\Delta \ln \text{Trade Weighted Dollar Index}$ ) and the NSE NIFTY ( $\Delta \ln \text{NSE}$ ) was further examined to confirm the impact of an improved trade relationship between India and the US on NSE NIFTY. The VAR Granger test results with an F value of 4.056 is significant at 0.012 percent significance

level and further strengthens the causal relationship. However, the reverse relationship between the variables did not seem to be significant, with an F value of 2.3056 and p-value of 0.129. Table 7 shows the results of the Granger causality test for the period January 2005 to December 2018.

**Table 7.** Granger causality in VAR Framework for the period January 2005–December 2018.

Independent Variables	Dependent Variable	F Statistic	p-Value	Independent Variable	Dependent Variable	F Statistic	p-Value
$\Delta \ln \text{USD}$	$\Delta \ln \text{NSE}$	6.990	0.000	$\Delta \ln \text{NSE}$	$\Delta \ln \text{USD}$	3.152	0.007
$\Delta \ln \text{EUR}$		1.0244	0.311		$\Delta \ln \text{EUR}$	2.044	0.152
$\Delta \ln \text{CNY}$		6.242	0.000		$\Delta \ln \text{CNY}$	1.7041	0.139
$\Delta \ln \text{JPY}$		10.847	$1.982 \times 10^{-5}$		$\Delta \ln \text{JPY}$	1.3125	0.269
$\Delta \ln \text{GBP}$		0.0057	0.939		$\Delta \ln \text{GBP}$	0.6725	0.644

Note: Granger Causality Null—No Granger Causality.

For the same period, the trade with Japan totaled \$17.63 billion ([Embassy of India Japan 2020](#)). The recent currency swap worth \$75 billion between India and Japan might have had a short-term impact on the capital markets in the country, especially on the National Stock Exchange ([Oberoi 2018](#)). Moreover, Japanese Yen being a safe haven, will appreciate at times of financial turbulence, hence, it can be said that Yen impacted the movements of the Nifty 50 index, i.e., when Yen appreciates it is possible that there is a pessimism in the global markets ([Hampson n.d.](#)) and this would eventually pass on to the Indian index. The analysis also points out that the INR market has started tracking the Chinese Yuan. In fact, Indian Rupee is more correlated with the Chinese Yuan than the USD. The reason being the rapidly expanding money supply pressurizing the Chinese regulators to devalue their currency, which makes the exports competitive for China. This would adversely impact the Indian trade, which gets reflected in the Nifty 50 index ([Patel n.d.](#)). Based on the results, Hypotheses 2, 4, and 5 are accepted.

It is also interesting to observe that the causal relationship between EUR, GBP, and NSE NIFTY index is not significant because the FII equity inflows in terms of GBP and Euro are not as massive as USD and JPY. The fund flows in terms of these currency denominations suggest that they are insufficient to cause any significant change in the Nifty 50 index ([NSDL n.d.](#)).

The Bidirectional causality results further revealed that the Nifty 50 index does influence the USD/INR exchange rates. There are multiple reasons for this. The prominent one being the Foreign Institutional Investors (FII). When the FIIs pour in money to the Indian stock markets, the money, in terms of USD, needs to be converted to INR, creating a demand for INR and its appreciation against USD ([Khatri n.d.](#)). From the financial year 2011–2012, the United States of America has the most FII investments in equity flows to the Indian markets. The USA's FPI inflow towards the Indian Equity markets for 2017–2018 was Rs.9,65,768 crores ([NSDL n.d.](#)). These flows substantiate that a part of USD/INR exchange rate fluctuations is due to dollar-denominated equity flows into the Indian Market. A stronger Nifty 50 would mean that the rupee stays strong against dollars because of the dollar infusion into the Indian equity market.

#### 4.5. Impulse Response Functions

Impulse Response functions are used to figure out to what extent a shock in one of the exchange rates under study would affect the NSE NIFTY index over a period of 12 months for the entire study period. The results are given in Figure 3.

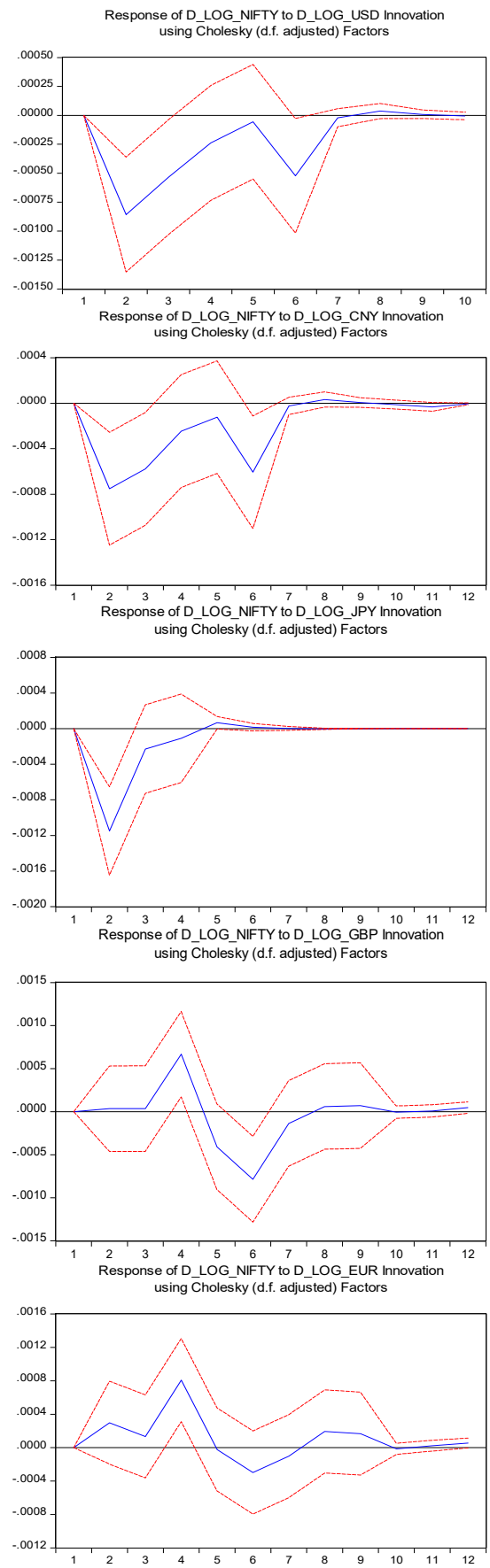


Figure 3. Impulse Response Analysis Results.

The response of D\_LOG\_NIFTY to a shock caused by D\_LOG\_USD shows that NIFTY falls in the first and second period and then rises and converge to zero close to the 10th period. In the case of Chinese Yuan, a shock caused by D\_LOG\_CNY will result in the drop of D\_LOG\_NIFTY in the first two periods and then rises and converge to zero close to 10th period. D\_LOG\_NIFTY responds to the shock caused by D\_LOG\_JPY in a similar way as it responded to D\_LOG\_USD and D\_LOG\_CNY, where NIFTY drops starkly in the first two periods and then rises. Unlike the other two currencies, NIFTY totally converges to zero after the seventh period. The shock caused by D\_LOG\_GBP and D\_LOG\_EUR on D\_LOG\_NIFTY shows a different trend from the first three cases. D\_LOG\_NIFTY rises from second period and then falls starkly until sixth period. The shock caused by D\_LOG\_EUR also follows a similar pattern. However, both D\_LOG\_GBP and D\_LOG\_EUR do not converge to zero like the other three currencies.

## 5. Conclusions

This study aimed at analyzing the short-run and long-run linkages between exchange rates and NSE NIFTY index through a Johansen cointegration and a Granger causality test. The results of the cointegration analysis show that no long-run relationship exists between exchange rates and NIFTY index. This result implies that the variations in the exchange rates under this study may not have a significant influence on NIFTY index in the long-run, hence, the investors who are into long-term trading do not necessarily have to be cautious about the short-term fluctuations in the exchange rates. The Granger causality test results indicate that the USD, CNY, JPY exchange rates have a positive and significant relationship with NIFTY index. The findings are further supported by the Impulse Response Function analysis that gives the time required for convergence of NIFTY index following an initial shock. This finding indicates that the stakeholders need to keenly follow the short-run fluctuations in the above-mentioned exchange rates if their intention is to engage in short-term trading. One of the limitations of this study is that it uses conventional methods of cointegration and Granger causality to assess the relationship between the variables. Future studies in this area should consider the volatility spillover effects using more appropriate methods.

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