

Mukhopadhyay, Debabrata; Sarkar, Nityananda

## Article

# Demonetization and Its Effects on BSE SENSEX and Some Sectoral Indices: An Exploratory Econometric Analysis

International Econometric Review (IER)

## Provided in Cooperation with:

Econometric Research Association (ERA), Ankara

*Suggested Citation:* Mukhopadhyay, Debabrata; Sarkar, Nityananda (2019) : Demonetization and Its Effects on BSE SENSEX and Some Sectoral Indices: An Exploratory Econometric Analysis, International Econometric Review (IER), ISSN 1308-8815, Econometric Research Association (ERA), Ankara, Vol. 11, Iss. 2, pp. 38-57,  
<https://doi.org/10.33818/ier.473544>

This Version is available at:

<https://hdl.handle.net/10419/238833>

## Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

## Terms of use:

*Documents in EconStor may be saved and copied for your personal and scholarly purposes.*

*You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.*

*If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.*



<https://creativecommons.org/licenses/by-nc-nd/4.0/>

## **Demonetization and Its Effects on BSE SENSEX and Some Sectoral Indices: An Exploratory Econometric Analysis**

**Debabrata Mukhopadhyay<sup>®</sup> and Nityananda Sarkar**

West Bengal State University and Indian Statistical Institute

Received: 13.02.2018    Accepted: 07.05.2019    Published: 30.09.2019  
doi: 10.33818/ier.473544

### **ABSTRACT**

This study has carried out some preliminary time series analyses to examine the impacts of demonetization which was carried out in India on 8 November 2016, on the well-known Indian stock index, BSE SENSEX, and four major sectoral indices viz., BSE BANKEX, BSE Auto, BSE Reality and BSE Smallcap, using daily level time series data covering the period 1 January 2016 to 31 May 2017. Apart from examining the stationarity/nonstationarity property and existence of structural breaks after demonetization in these series, the paper has also studied the trend behavior and returns models for both the pre- and post-demonetization periods. This study has found that while there is more than one break in all the five series at their level values, there is only one structural break after demonetization. It has been found that the trend function for all the index series broadly gives support to this finding of one break after demonetization. Further, some changes have been observed in the stationary models for the two sub-periods of pre- and post-demonetization for all but BSE Auto index. Finally, except for BSE SENSEX, no change in the status of stationarity/nonstationarity in the pre- and post- demonetization periods has been found in the other series.

**Key words:** *Bai-Perron Multiple Break Points Test, Demonetization, Stationarity, Structural Breaks, Sup MZ Test*

JEL Classifications: C22, G18, G28

## **1. INTRODUCTION**

Demonetization can be described as a monetary action where the monetary authority withdraws currency notes of higher denominations from circulation and replaces the same with new bank notes. This is a very rare monetary policy decision taken by any government to tackle extreme economic situations like hyperinflation, severe corruption, and intolerable size of parallel economy. In this context, mention may be made of a paper by Rogoff (2014) where he has advocated in favour of phasing out paper currency for a legal economy while discussing the costs and benefits of such a step by any central bank. It is also mentioned in the literature that paper currency makes transactions anonymous which is an essential property of money, giving it a certain form of anonymity (see, for details; Kiyotaki and Wright, 1989). In most countries, a large percentage of currency (over 50%) is used to hide transactions because of anonymity character of paper currency (see, for instance; Rogoff, 1998, 2002). In a recent book, Rogoff

---

<sup>®</sup> Debabrata Mukhopadhyay, West Bengal State University, Barasat, North 24 Parganas, Kolkata 700 126, India (email: [debu1641975@yahoo.co.in](mailto:debu1641975@yahoo.co.in)).

Nityananda Sarkar, Indian Statistical Institute, 203 B.T. Road, Kolkata 700 108, India (email: [tanya@isical.ac.in](mailto:tanya@isical.ac.in)). The authors would like to thank sincerely the two anonymous referees and the Editor for their careful reading of the manuscript, which has led to considerable improvement in the final version of the paper.

(2016) has argued how paper money can also cripple monetary policy as it does not allow interest rate to go below zero.

The Government of India went for a massive demonetization exercise on 8 November 2016, by withdrawing the legal tender status of the then existing Rupees (Rs.)<sup>1</sup> 500/- and Rs. 1000/- currency notes which constituted around 86% of the total currency in circulation as per the Reserve Bank of India's (RBI)<sup>2</sup> Database on Indian Economy. According to the announcement made by the government, this rare monetary action was taken in order to contain the rising incidence of fake notes and black money in the country leading to running of a huge and ever increasing parallel economy within the country. In an affidavit submitted to the Supreme Court by the Government of India, it has also been mentioned that "...the withdrawal of existing high denomination banknotes will curb funding of terrorists with the proceeds of fake Indian currency notes (FICN) and use of existing FICN network for subversive activities." A movement towards digital economy has also been advocated as a major argument in favour of demonetization. The government tends to think, based on some economic reasoning, that a developing economy should have lesser use of paper currency and higher technology-based transactions.

However, this demonetization decision came under severe criticism from some noted economists on the ground that (i) the move was authoritarian in nature, (ii) it is very likely to lead to immense difficulties to common people, and (iii) it would adversely affect business and commerce, especially in the small and medium sectors, and consequently lead to shrinkage of production and employment in those sectors. Some also criticized it in terms of its effectiveness to significantly reduce 'black money', which is a process outcome of the prevailing economic system rather than a one-time phenomenon, especially in a vast and not-so-efficiently managed developing country like India.

Although this is not the first time that demonetization has been undertaken in India<sup>3</sup>, but the unanticipated manner in which it happened and the unpreparedness that the country exhibited made it a major public debate issue for quite some time. Also, unlike the earlier two occasions in 1946 and 1978, this has affected a very large section of the population. At the world level, only very few countries ever undertook such a monetary action. Even in that, most of these countries that followed demonetization, failed miserably in terms of achieving their goals. Among the least developed economies that carried out the rare monetary policy of demonetization, most were Latin American countries, and their objective was to control hyperinflation and/or high level of corruption. However, in order to understand the implication of this major policy decision in a big and populated country like India, it is noteworthy to consider some intrinsic features of the Indian economy which are relevant in this context.

Indian economy is predominantly a cash-intensive economy. According to the 2013 report of the Institute for Business in the Global Context (IBGC) on cash outlook, it has been noted that the value of notes and coins in circulation as a percentage of GDP was 12.04% for India, which is far higher than those for countries like Brazil (3.93%), Mexico (5.32%) and South Africa (3.72%). The ratio<sup>4</sup> of  $M0/M2$  is also high in India compared to other comparable developing countries, with almost 87% transactions in 2012 being cash-based transactions. According to

---

<sup>1</sup> The currency note of India is called Rupee or, Re. in abbreviation; Rs. stands for abbreviation of Rupees.

<sup>2</sup> The Central Bank in India is called the Reserve Bank of India.

<sup>3</sup> Currency notes of higher denominations were earlier withdrawn from circulation in India in 1946 and 1978, but those affected only the privileged few.

<sup>4</sup>  $M0$  and  $M2$  stand for narrow and broad money supply, respectively.

the Global Findex data of the World Bank, 53% Indians over age of 15 years had bank accounts in 2014 although India also has the highest number of dormant bank accounts. Another crucial feature of the Indian economy is that life of the economy is still in the informal sector which, by the very nature of its activities, is highly cash dominated. The latest Economic Census of 2013-14 has reported that 98.6% establishments in India are in the informal sector. Further, employment figures show that 82% of total workers are in the informal sector (of which 52% are self-employed and 30% casual workers) and 10% informal workers are in the formal sector. A mere 8% of the workers are engaged in the formal sector as formal workers.

The consumption-production linkage between informal and formal sectors that exists in a dual economy like India is also important in this context. The demonetization process which may severely affect the real economic activity of cash-intensive informal economy is also channelized to the formal sector and to the financial sector, especially to the stock market because of interdependences between real and financial sectors. In order to determine the macroeconomic impact of this unanticipated monetary shock, one can find the contraction of aggregate demand due to cash crunch and real balance effect i.e., the Pigou effect. The aggregate supply curve may also move left due to contraction of money supply ( $M0$ ) as a result of this demonetization. Thus, real economic activity may be reduced. The effect on inflation is; however, ambiguous, being dependent on the relative strength of change in aggregate demand and supply. Demonetization also causes disproportionate effect on various sectors of the economy. The work done by National Institute of Public Finance and Policy (NIPFP) in November 2016 (Rao et al., 2016), has clearly pointed out that the sectors to be adversely affected are all those sectors where demand is usually backed by cash like, for example, transport services, real estate, and small business. Another sector that is crucially dependent on cash for carrying out its activities and hence likely to be affected by demonetization is the banking sector.

Considering the issues discussed above, the basic purpose of this work is to study the impacts of this demonetization, if any, on some important sectors of Indian economy which are largely dependent on cash transactions. But lack of availability of sufficient data is a practical constraint for doing such a study since for most of the relevant macroeconomic variables, the data are available at low frequency i.e., either at monthly or quarterly level making the number of sample observations in the post-demonetization period to be too few to carry out any meaningful time series analysis. Keeping this in mind, we have confined our attention to Indian stock market only where daily level data are available. This issue apart, it may be relevant to ask why it is worth studying the impact of demonetization which is a rare monetary event in the history of any economy, on some stock indices. It is very well recognized that stock market is a barometer of the economy where the effects of major economic policy decisions are expected to be instantaneously transmitted to the agents' expectation formation about the future state of the economy, which is more commonly understood through the stock market. Ioannidis and Kontonikas (2006) pointed out that stock prices are commonly regarded as being highly sensitive to economic conditions. In the context of transmission mechanism through the stock market, monetary policy actions affect stock prices, which themselves are linked to the real economy through their influence on consumption spending i.e., the wealth effect channel, and investment spending i.e., the balance sheet channel. Thus, impact of this massive monetary action may have some bearing in the capital market, especially in the stock market of the country.

In this limited study, we have chosen five indices which, we thought, are likely to be affected by demonetization. Thus, to be specific, this paper empirically studies the possible effects of

demonetization on four major sectoral indices of the Bombay Stock Exchange (BSE)<sup>5</sup>; namely, S&P BSE BANKEX, S&P BSE Auto, S&P BSE Reality, and S&P BSE Smallcap. Apart from these indices of four sectors where transactions are heavily cash dependent, we have also taken the most widely quoted index of the Indian stock market *viz.*, S&P BSE SENSEX. The reason for including the last one is because of its importance as an overall index of Indian stock prices, and the fact that nothing, as such, is known about how the likely impacts of demonetization on different sectors may affect this widely-quoted overall stock index of India through different transmission channels.

The data refer to the daily level closing prices of the five indices over the period 1 January 2016 to 31 May 2017. The purpose for choosing the time series covering a period prior to demonetization is to be able to compare the results obtained, based on analysis of time series, for the post-demonetization period with those of the pre-demonetization period, wherever such comparisons are meaningful and relevant.

The aim of this paper is very limited *viz.*, this study is a preliminary analysis of each of the five indices of the Indian stock market, as mentioned above, based on application of econometric tools available for univariate time series analysis. The sole aim is to find if there has been any (statistically) significant change/effect on these series, both at level and difference values, as a result of this very important monetary action by the Government of India. To the best of our knowledge, no such study has yet been done. What are available are three studies on consequences of demonetization based on economic theoretic reasoning, and some newspaper articles. The first such study is due to Rao et al. (2016) where they have pointed out that in the long run the benefits of demonetization will crucially depend upon the availability of credit, spending, level of activity and government finances. The second paper by Dasgupta (2016) has questioned, following Robert Lucas's Nobel lecture, the merits of economic policies that assume the form of random shocks to an economic system while explaining the theoretical issues of demonetization. The last study by Muthulakshmi and Kalaimani (2016) on this issue basically discusses the effectiveness of demonetization decision to eradicate parallel economy.

The paper is organised as follows. Section 2 describes the data and methodology. Empirical findings are discussed in Section 3. Conclusions are made in Section 4.

## **2. DATA AND METHODOLOGY**

The study involves time series data of five indices of the Indian stock market<sup>6</sup> *viz.*, BSE SENSEX, BSE BANKEX, BSE Auto, BSE Reality and BSE Smallcap at daily closing prices covering the period 1 January 2016 to 31 May 2017. There are, in all, 349 observations for each series. All the time series have been downloaded from [www.bseindia.com](http://www.bseindia.com). The BSE SENSEX is the most representative stock index of India comprising 30 prominent stocks traded actively in the exchange and it covers all the key sectors of the economy. The other four indices *viz.*, BSE BANKEX, BSE Auto, BSE Reality and BSE Smallcap are sectoral indices.

The BSE BANKEX consists of the companies classified as banks in the BSE 500 index which are found to be good in terms of trading frequency and free float market capitalization. This index comprises both private sector banks such as ICICI, and HDFC, and public sector banks like the State Bank of India, and Bank of Baroda. The sectoral auto index consists of the

---

<sup>5</sup> The Bombay Stock Exchange which was established in 1875 is the oldest stock exchange of India.

<sup>6</sup> All these indices are now prefaced with S&P; for example, BSE SENSEX is now called as S&P BSE SENSEX. However, for the sake of convenience, we refer these indices, all throughout the paper, without this preface 'S&P'.



automobile companies enlisted in BSE 500 index, and includes major players in the market such as Tata Motors, Mahindra & Mahindra, Bajaj Auto, and Maruti Suzuki. The index that represents real estate sector is the BSE Realty consisting of the real estate companies such as DLF Ltd., and Unitech Ltd. which are listed in BSE 500 index. The last sectoral index considered is BSE Smallcap that includes companies which are small in terms of market capitalization and are not very well known. The choice of the last two sectoral indices has been primarily guided by consideration of the likelihood of their being affected by demonetization because of the nature of economic activities represented through stock prices of these indices.

In what follows next, we first describe the methodological approach that has been followed in this paper. As already stated, an econometric analysis of these five time series at the univariate level has been done in this paper. To be specific, we have first studied two most important and fundamental properties of each series *viz.*, stationarity and structural break. Thereafter, we have obtained the ‘best’ univariate model for the stationary series, and finally, carried out trend analysis in case a series is found to be nonstationary. These have been done for the full sample as well as for the relevant sub-periods of pre-demonetization and post-demonetization as thought to be appropriate and relevant. The core point in all these exercises is to find if there has been any impact in the important time series properties as well as the data generating process of these five series due to this rare monetary decision of demonetization.

The stationarity property has been checked by applying the well-known augmented Dickey Fuller (ADF) test, both at level values as well as differences. All the indices have been taken at their natural logarithmic values. Thus, the values at first difference stand for the continuously compounded rate of returns on the corresponding index.

Next, we have performed a test for the presence of any structural break(s), if any, in each of these five series. This exercise has been carried out both at level values where the series is non stationary in the sense of having trend, and (first) difference values where the series is stationary. The existence of breaks in the series has been studied both at level and difference values since structural breaks may occur at either or both cases/situations. It is probable that such an unanticipated monetary shock in terms of this massive demonetization action would lead to permanent change(s)/breaks(s) in the data generating process of the time series either or both at nonstationary and stationary levels. The literature on structural break(s) starting with the classic work by Chow (1960) is vast and comprehensive. The Chow test has few well-known limitations. Andrews in his seminal work in 1993 proposed a proper statistical test for a stationary series with an unknown single structural break point. Further, a new structural break test, called the *MZ* test that tests simultaneously for break in regression coefficients as well as in variance, was developed by Maasoumi et al. (2010) in 2010. Subsequently, *sup MZ* test was proposed by Ahmed et al. (2017), which caters to an unknown breakpoint, and comparison of this test was made with the *sup F*-test of Andrews (1993). We have used the *sup MZ* test to check whether our finding, based on the Bai-Perron test, on structural break in each series occurring after demonetization, is supported or not by the *sup MZ* test.

As regards testing for the presence of multiple structural breaks and that too in case of nonstationary series as well, the first major breakthrough was given by Bai and Perron (1998, 2001). In a subsequent paper in 2003, they advocated, based on extensive simulations, that under very general conditions on the nature of data and the error term, the following testing procedure involving basically two test statistics *viz.*, the *UD max* test and/or the *WD max* test, and the *sup F*-type test i.e., a sequential test of the null hypothesis of  $l$  breaks *versus* the alternative of  $(l+1)$  breaks, be followed. First, the *UD max* and/or *WD max* tests are used to

see if at least one break is present. If these indicate the presence of at least one break, then the number of breaks can be decided based upon a sequential examination of the *sup F*-type statistic. It may be noted that in the later, the test  $F(1|0)$  is ignored.

The Bai-Perron multiple breaks point test has been applied to find structural breaks in all the five series. In case the test suggests the existence of one or more structural breaks, then appropriate models have been obtained for the five series at stationary level i.e., return series, for both the pre- and post-demonetization periods in order to find if there is any significant difference in the two models for these two periods.

As already stated, we have also applied the *sup MZ* test to find if the estimated break date after demonetization is found to be the same by the *sup MZ* test for each of the five series. To be explicit, what we have done is that if the Bai-Perron test has identified, say, 5 break points  $T_1, T_2, T_3, T_4$  and  $T_5$  of which  $T_3$  refers to the estimated break point after demonetization, then we have used the data covering the period from  $T_2$  to  $T_4$  to re-test the existence of a break point during this period by applying the *sup MZ* test.

In case trend is found to exist in any time series, trend analysis has been carried out in order to find if demonetization has led to any change in the trend function. To that end, the following trend models for the logarithmic price series<sup>7</sup>  $p_{it} = \ln(P_{it})$ ,  $i=1,2,3,4,5$  where  $P_{1t}, P_{2t}, P_{3t}, P_{4t}$  and  $P_{5t}$  stand for the daily closing prices of BSE SENSEX, BSE BANKEX, BSE Auto, BSE Reality, and BSE Smallcap, respectively, have been applied. For considering the break date which is taken to be given in trend analysis, we have used the estimated break date(s) obtained for the corresponding series using the Bai-Perron test. The five trend models considered are stated below.

$$\text{Model A: } p_{it} = \alpha_{0A}^i + \alpha_{1A}^i D_{1t} + \beta_{1A}^i t + \varepsilon_{tA}^i$$

where  $D_{1t} = 1$ , if  $t > T_B$

= 0, otherwise

and  $T_B$  is the break date. The subscript/superscript  $i$  stands for the  $i$ -th index with  $i=1,2,3,4,5$ .

$$\text{Model B: } p_{it} = \alpha_{0B}^i + \alpha_{2B}^i D_{2t} + \beta_{1B}^i t + \varepsilon_{tB}^i$$

where  $D_{2t} = 1$ , if  $T_B < t < T_C$

= 0, otherwise

and  $T_C$  is the last time point (i.e., day) upto when the effect of demonetization, if any, remains.

$$\text{Model C: } p_{it} = \alpha_{0C}^i + \beta_{1C}^i t + \beta_{2C}^i D_{3t} + \varepsilon_{tC}^i$$

where  $D_{3t} = t - T_B$ , if  $t > T_B$

= 0, otherwise.

$$\text{Model D: } p_{it} = \alpha_{0D}^i + \alpha_{4D}^i D_{1t} + \beta_{1D}^i t + \beta_{2D}^i D_{3t} + \varepsilon_{tD}^i.$$

---

<sup>7</sup> Since, as already stated, all the five series have been transformed to their natural logarithmic values so that the difference represents the continuously compounded rate of returns, the trend analysis has been carried out on log-transformed series, and the results presented and discussed thereof. However, trend analysis was also done on level values without log-transformation i.e.,  $P_{it}$ . But the conclusions were found to be the same as with the log-transformed values.

It may be mentioned at this stage that we are using different trend models instead of a single model with the idea that effect of structural break(s) may vary across intercept and slope parameters differently as represented in Models A through D. It may be summarized that ‘Model A’ is a one-time change in the intercept of the trend function where the dummy variable is a step dummy. ‘Model B’ represents the same but with an extended pulse dummy. It may be noted that the model with such dummy variables enables us to know at what point of time demonetization has made impact, if any, on stock indices, and also up to what time point. ‘Model C’ allows for a change in the slope of the trend function without any sudden change in the level at the time of structural break. Finally, ‘Model D’ considers both changes in intercept and slope of the trend function.

In a seminal work Nelson and Plosser (1982) argued that if a time series is trend stationary (TS) then the effects of random shocks disappear over time i.e., the shocks are transitory in nature, whereas for a difference stationary (DS) process, random shocks have permanent effects. Since demonetization in India was an unanticipated random shock, it is meaningful to find if these shocks in the post-demonetization period are temporary or permanent in nature. With this view in mind, we checked the stationarity status of each index at logarithmic level values separately for the pre- and post-demonetization periods by applying the ADF test. It is well known that the ADF test is also used for finding if a time series is DS or TS in case it is found to be nonstationary.

All the computations have been carried out using the software EViews 9.

### 3. EMPIRICAL FINDINGS

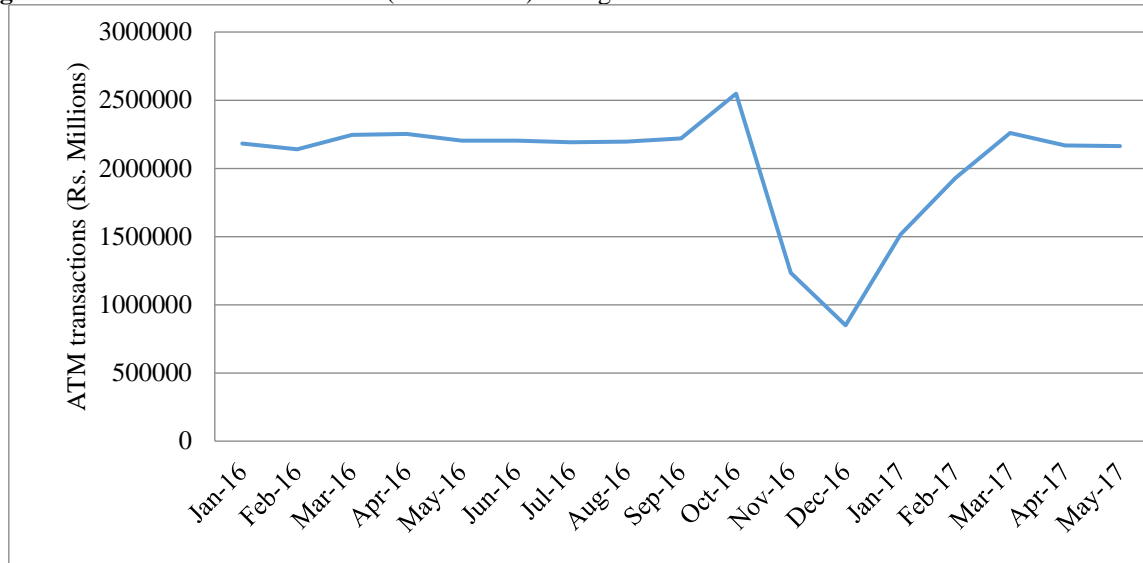
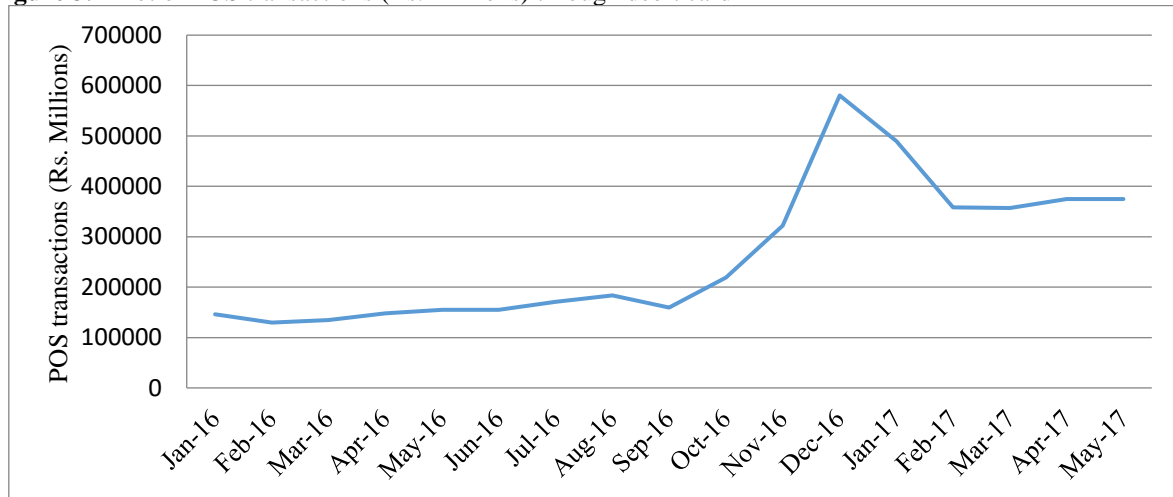
In this section, we discuss the empirical results following the econometric methodology presented in the preceding section. We first present, in Table 3.1, the summary statistics of all the five series at their first difference (log) level values i.e., for the cumulated returns series. It is observed from Table 3.1 that all the mean returns are positive. Further, among the five mean returns, the one on BSE Reality index is the highest, and that on BSE SENSEX is the lowest. Average annual returns values vary from 12% to 24%. We also find that standard deviation value is also the highest for returns on BSE Reality and lowest in case of BSE SENSEX. All the series except BSE BANKEX have negative skewness values. Further, all the series have excess kurtosis values i.e., all the return distributions are leptokurtic.

<i>Returns series</i>	Mean	Median	Std. deviation	Skewness	Kurtosis
<i>BSE SENSEX</i>	0.0005	0.0005	0.008	-0.093	4.499
<i>BSE BANKEX</i>	0.0009	0.0009	0.012	0.147	4.178
<i>BSE Auto</i>	0.0007	0.0016	0.012	-0.320	5.073
<i>BSE Reality</i>	0.001	0.002	0.019	-0.647	7.429
<i>BSE Smallcap</i>	0.0008	0.002	0.011	-1.148	6.466

**Table 3.1** Summary Statistics of Returns on the Five Indices for Full Sample.

Before proceeding further, we take note of an observation made from visual inspection of the plots showing bank ATM transactions and Point of Sale (POS) transactions at monthly level for the period under consideration. From these two plots given in Figures 3.1 and 3.2, we note that the ATM transaction decreased sharply during the three months (including November) after demonetization and POS transaction increased substantially during the same period.

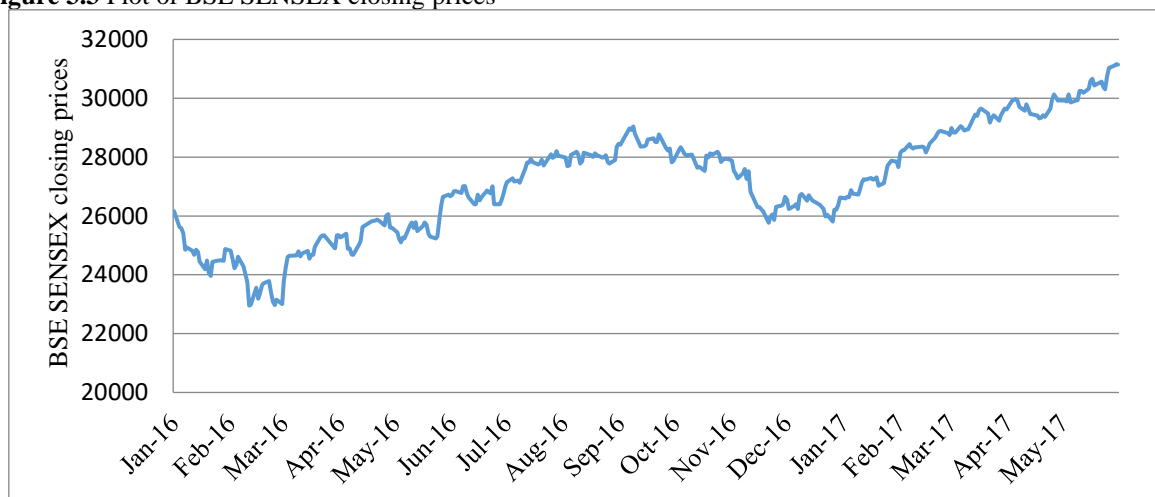


**Figure 3.1** Plot of ATM Transactions (Rs. Millions) through debit card**Figure 3.2** Plot of POS transactions (Rs. Millions) through debit card

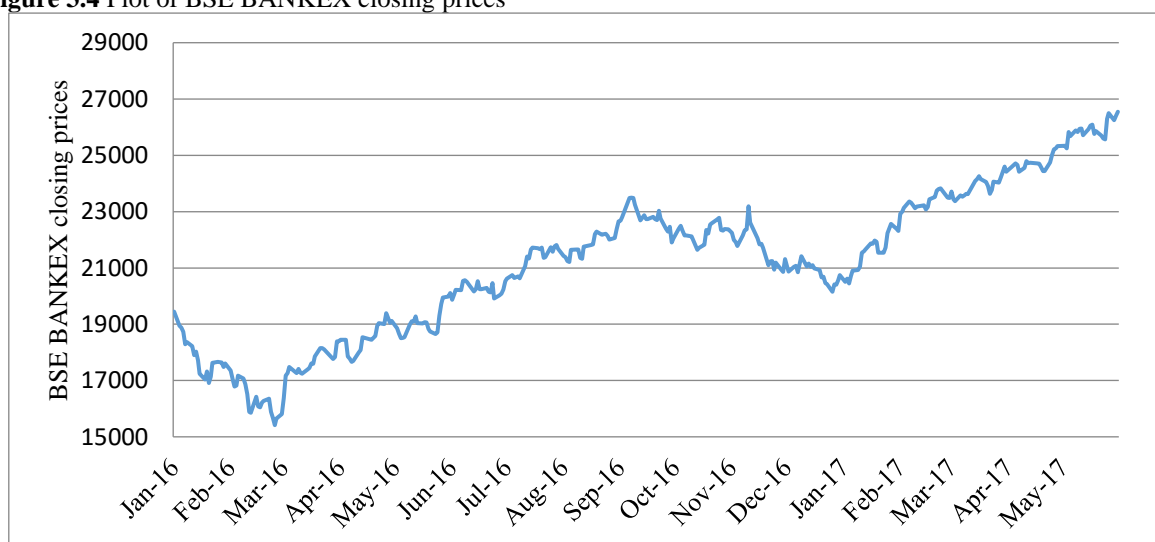
However, although ATM transactions was back to the pre-demonetization level by March, 2017, POS transactions even though decreased from the peak achieved in January 2017, remained substantially higher than that of the pre-demonetization period indicating thereby a structural shift in the movement towards digital transactions in the sample period. To figure out visually the likely changes in the stock index series after demonetization, we have given diagrammatic representations of all the five series at level values. In Figure 3.3, we have plotted the daily closing price for BSE SENSEX. Likewise, Figures 3.4 through 3.7 represent the same for BSE BANKEX, BSE Auto, BSE Reality and BSE Smallcap series, respectively.

It is evident from the plots that right after the date of demonetization i.e., 8 November 2016, there is a slowly declining pattern till about middle of January 2017, for all the five index series, but thereafter all the series have started rising. It is also noted from these plots that there is no exhibited permanent downward or upward shift after demonetization in the sample period for any of the five stock indices. At this stage, it may be worth noting that soon after demonetization there was gradual remonetisation i.e., replacement of old discontinued currency notes with new ones in the economy, and this might have contributed to all the five series getting back to square one after some lags.

**Figure 3.3** Plot of BSE SENSEX closing prices

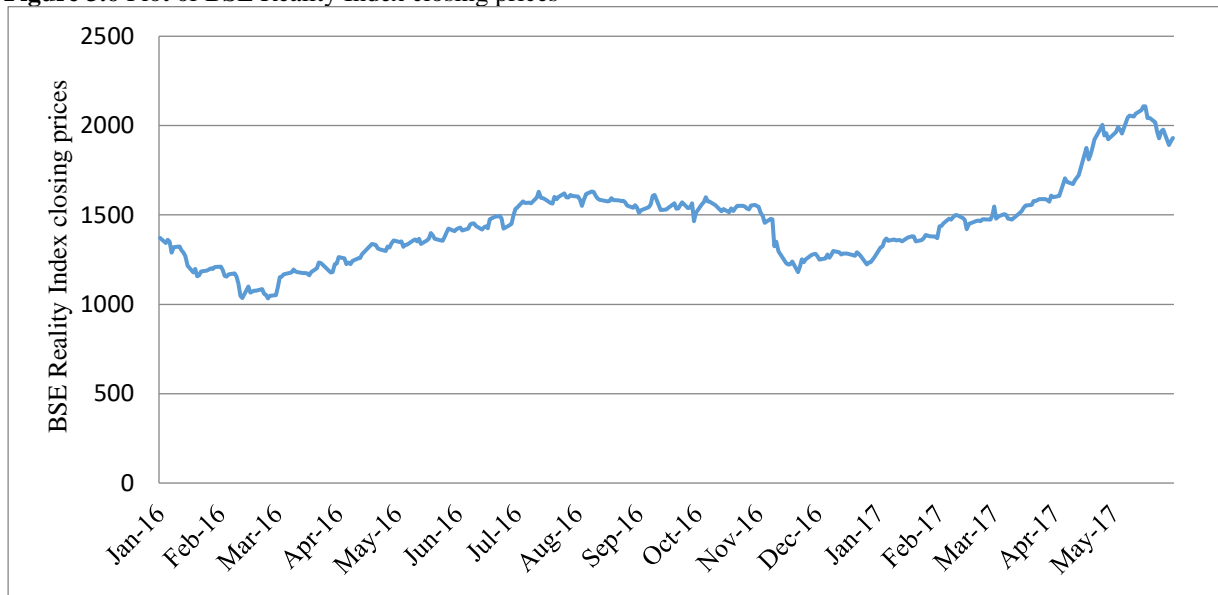
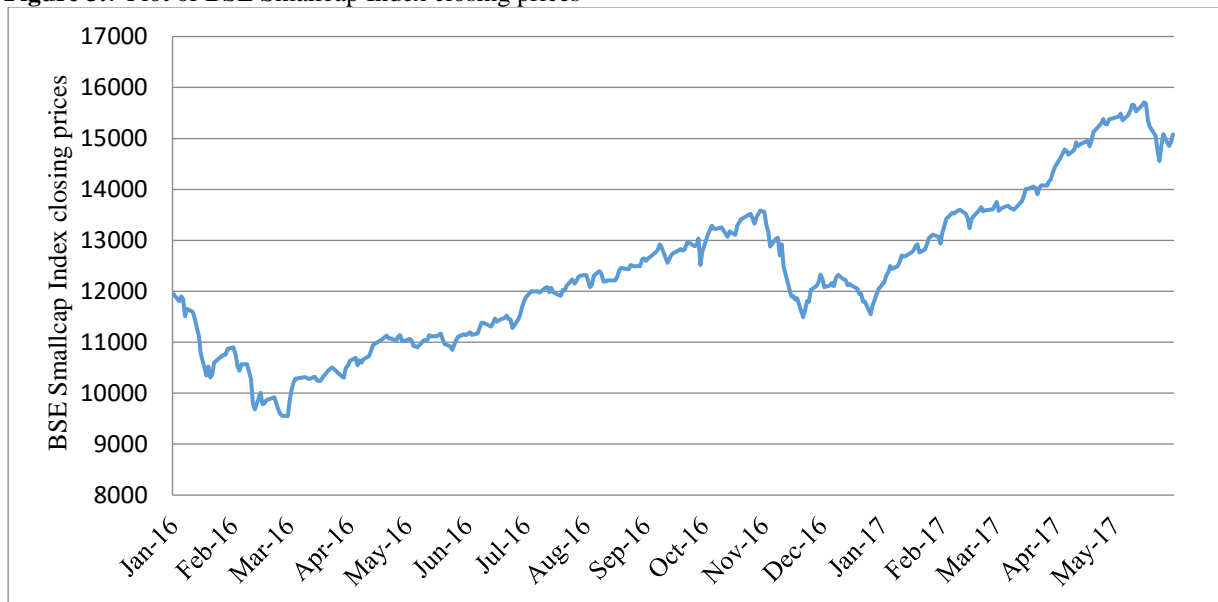


**Figure 3.4** Plot of BSE BANKEX closing prices



**Figure 3.5** Plot of BSE Auto Index closing prices



**Figure 3.6** Plot of BSE Reality Index closing prices**Figure 3.7** Plot of BSE Smallcap Index closing prices

### 3.1. Results on Stationarity and Structural Breaks (Based On Full Sample)

The results on the ADF tests on all the five series are presented in Table 3.2 below. All the five indices at their level (in natural logarithmic) values have been found to be nonstationary since the null of unit root could not be rejected for all these series at 5% level of significance. Thereafter we tested if these series are stationary at first difference level. It is evident from the ADF test statistics values that all are significant at 1% level, showing thereby that there is no unit root in the first difference series. This also implies that all the five indices are integrated of order 1. We further note that in the deterministic part of the ADF estimating equation, only the intercept is found to be significant at first difference level, which suggests that deterministic trend is significant at level values for all the series.

<i>Index</i>	ADF test statistic value		<i>p</i> -value		Deterministic component	
	Level	First diff.	Level	First diff.	Level	First diff.
<i>BSE SENSEX</i>	-2.407	-17.497	0.375	0.000	constant, trend	constant
<i>BSE BANKEX</i>	-2.651	-16.974	0.258	0.000	constant, trend	constant
<i>BSE Auto</i>	-2.336	-17.704	0.413	0.000	constant, trend	constant
<i>BSE Reality</i>	-1.630	-18.509	0.772	0.000	trend	constant
<i>BSE Smallcap</i>	-3.270	-16.277	0.073	0.000	constant, trend	constant

**Table 3.2** Results of Unit Root Test based on Full Sample.

Notes: All the test statistic values are compared with MacKinnon (1996) one-sided critical values. Level values are in natural logarithm.

Next, we report the results on the impact of demonetization in structural change in these series at both nonstationary and stationary levels i.e., for both the index series (i.e., at level values) and the returns series (i.e., first difference values), by following the multiple structural break points test proposed by Bai and Perron (1998, 2001, 2003), as discussed in the preceding section. We have presented the findings of this test in Tables 3.3 and 3.4 along with the (estimated) break dates.

<i>Index</i>	$UD_{max}$ statistic value		$WD_{max}$ statistic value	
	Level	First diff.	Level	First diff.
<i>BSE SENSEX</i>	560.775*	10.286	724.505*	10.286
<i>BSE BANKEX</i>	538.127*	14.361*	738.315*	14.361*
<i>BSE Auto</i>	833.264*	9.987	875.635*	12.02
<i>BSE Reality</i>	828.297*	9.116	1321.358*	11.872
<i>BSE Smallcap</i>	638.739*	12.969*	751.419*	14.018*

**Table 3.3** Results of  $UD_{max}$  and  $WD_{max}$  Tests based on Full Sample.

Note: \* indicates significant value at 5% level of significance.

<i>Index</i>	Level		First Difference	
	Total number of break dates	Estimated break date after demonetization	Total number of break dates	Estimated break date
<i>BSE SENSEX</i>	4	11 November 2016	0	-
<i>BSE BANKEX</i>	4	16 November 2016	1	22 March 2016 <sup>#</sup>
<i>BSE Auto</i>	5	11 November 2016	0	-
<i>BSE Reality</i>	4	9 November 2016	0	-
<i>BSE Smallcap</i>	4	11 November 2016	1	22 April 2016 <sup>#</sup>

**Table 3.4** Results of the Bai-Perron Multiple Structural Break Points Test based on Full Sample.

Notes: Trimming parameter value is 0.15; level of significance considered is 0.05.

# indicates that both these dates are before demonetization.

By applying the  $UD_{max}$  and  $WD_{max}$  tests, it is found that the null hypothesis of ‘no break’ is rejected in favour of ‘one break’ for all the series at level values. However, for the difference series, the null of ‘no break’ could not be rejected for three series, namely, BSE SENSEX, BSE Auto and BSE Reality. But the  $UD_{max}$  and  $WD_{max}$  test statistics suggest that the null hypothesis of ‘no break’ is rejected for BSE BANKEX and BSE Smallcap at first difference values. Thereafter the sequential *sup F*-type test with trimming parameter value 0.15 was carried out, and the number of break dates in each series was obtained along with the estimated break dates. The results of this test with level values show that for all the indices except BSE Auto series, the number of breaks obtained in the entire series is four. For BSE auto series, the number of the estimated break dates has been found to be five.

It is also important to note that for each of the five indices, only one break in the series has been found immediately after demonetization. The estimated break dates for the five series at level values after demonetization are 11 November 2016, 16 November 2016, 11 November 2016, 9

November 2016, and 11 November 2016 for BSE SENSEX, BSE BANKEX, BSE Auto, BSE Reality, and BSE Smallcap, respectively. However, we have also presented the *sup F*-type test statistic values corresponding to each of the (estimated) break dates obtained for all the five stock prices at level values in order to empirically establish that there is strong evidence (global maximum) in favour of one of the break dates as compared to the other dated (local maxima). These test statistics values are presented in Table 3.5.

BSE SENSEX			BSE BANKEX		
Break No.	Estimated Break date	F-Statistic	Break No.	Estimated Break date	F-Statistic
Break date 1	11 November, 2016	238.839**	Break date 1	16, November, 2016	184.962**
Break date 2	11 July, 2016	54.758**	Break date 2	17 March, 2016	93.174**
Break date 3	17 March, 2016	64.280**	Break date 3	29 August, 2016	68.884**
Break date 4	25 January, 2017	21.352**	Break date 4	30 January, 2017	33.081**
BSE Auto			BSE Reality		
Break date 1	11 November, 2016	416.632**	Break date 1	9, November, 2016	398.666**
Break date 2	17 March, 2016	70.960**	Break date 2	21 March, 2016	70.139**
Break date 3	25 August, 2016	44.953**	Break date 3	4Julyt, 2016	98.142**
Break date 4	15 February, 2017	28.715**	Break date 4	16 March, 2017	71.693**
Break date 5	10 June,2016	9.064*			
BSE Smallcap					
Break date 1	15 November, 2016	159.957**	Break date 3	30 June, 2016	15.965*
Break date 2	17 March, 2016	246.923**	Break date 4	27 January, 2017	15.421*

**Table 3.5** Test Statistic values for all the break dates found from the Bai-Perron Test on level values

Note: \* and \*\* indicate significant values at 5% and 1% levels of significance, respectively. Only breaks at level values are reported.

We find from this table that the *F* statistic value is the highest corresponding to the post-demonetization break date for all but one (BSE Smallcap) series. This provides strong evidence for demonetization effect in terms of structural break on these stock indices. It may be stated that in the underlying model for each series, constant and linear trend have been taken as ‘breaking’ variables both at level and first difference values.

We now state the results of re-testing for the existence of structural break in the level values of all the five series after demonetization, by applying the *sup MZ* test. As stated earlier, the break point after demonetization considered for doing the re-testing by the *sup MZ* test is the one detected after demonetization by the Bai-Perron multiple break points test. The results of this re-test are presented in Table 3.6.

Index	Level		
	Estimated break date by the Bai-Perron test	Estimated break date by <i>sup MZ</i> test	<i>sup MZ</i> test statistic
BSE SENSEX	11 November 2016	10 November 2016	132.342*
BSE BANKEX	16 November 2016	21 December 2016	85.976*
BSE Auto	11 November 2016	10 November 2016	198.986*
BSE Reality	9 November 2016	8 November 2016	398.932*
BSE Smallcap	11 November 2016	3 November 2016	242.483*

**Table 3.6** Results of *sup MZ* Test for Re-testing the Existence of Break Point after Demonetization

Notes: \* indicates that the *sup MZ* test statistic value is significant at 1% level compared to tabulated chi-squared value of 9.21 with 2 degrees of freedom; Trimming parameter value is 0.15.

The existence of a structural break in the time period covering  $T_2$  to  $T_4$  is confirmed by the *sup MZ* test, as expectedly, for each of the series. Moreover, the results show that the estimated break date obtained by the *sup MZ* test is essentially the same - the difference being only of one day - as that determined by the Bai-Perron test tin case of three series *viz.*, BSE SENSEX, BSE



Auto and BSE Reality. For instance, the estimated break date for BSE SENSEX series is found to be 11 November 2016 by the Bai-Perron test while the re-test by *sup MZ* test yields this date to be 10 November 2016. Of the two remaining series, BSE Smallcap has a difference of 8 days which is once again close enough. It is only in case of BSE BANKEX that the two estimated break dates are found to be quite different with 16 November 2016 being the estimated break date by the Bai-Perron test, while it is 21 December 2016 by the *sup MZ* test.

Since estimated break date for BSE BANKEX has been found to be somewhat different by *sup MZ* test and the Bai-Perron test, an attempt was made to find which of the two tests estimates the unknown break points more accurately than the other. Any such exercise would require ensuring compatibility of the two tests in terms of their size performance. To that end, we first obtained a base null model which is the best fitting model with no breaks, and applied the Bai-Perron and *Sup MZ* tests at 5% level of significance. We then generated simulated data repeatedly and finally obtained the actual size of the two tests. It was, however, observed that the two tests were not compatible enough, and hence no further simulation-based exercise was done.

We have; however, tried to provide some explanations as to why the two tests might have picked up different estimated break dates for BSE BANKEX series from consideration of policies and actions initiated by the Reserve Bank of India (RBI) which is the Central Bank of India. The RBI issued a series of directives almost on daily basis since 8<sup>th</sup> December, 2016, the day of demonetization, on banking rules and regulations to implement demonetization. It periodically changed the rules on cash withdrawal and deposits at banks through the notifications. These notifications continued till 21 December when RBI finally removed restrictions on customers depositing old notes in excess of Rs 5,000. An RBI circular said banks would not be questioning customers depositing old notes if their accounts were compliant under ‘know your customer’ (KYC) norms. It may be observed from Figures 3.1 and 3.2 that ATM transactions improved from the third week of December, 2016. Therefore, it becomes obvious that all these directives made the banking sector more susceptible and this sector took much more time to stabilize compared to other sectors of the economy considered in our study. The likely departure from the assumptions underlying the two tests because of this kind of continuing changes in banking rules and regulations till the third week of December 2016 and hence the lack of stability in this particular time series may provide some explanations for the finding in case of BSE BANKEX.

Now, at the first difference level i.e., for return series, we note that the null hypothesis of ‘no break’ is rejected only for BSE BANKEX and BSE Smallcap. One structural break has been found in each of these two series, but neither break date is after demonetization. For the remaining three returns series there is no evidence of any structural break. Combining these with the earlier findings on the level values of the series, we thus conclude that these series have structural breaks in their level values but, barring two series, that too before demonetization, not at their first difference values.

Overall, it can be inferred that break or instability is a feature with all the five index series. But finding of one break date after 8 November 2016 also suggests that some impact of demonetization in respect of stability holds for each of the five series. Further, the presence of a number of breaks in each series at level values suggests that the trend breaks are temporary or ‘short-term’ in nature, and not permanent or ‘long term’, which broadly matches with the behavior exhibited in the plots. As regards to returns, we can conclude that, in respect of stability, all the five series have not been affected by demonetization in the same way since no

break has been found for BSE SENSEX, BSE Auto, and BSE Reality, and although one break has been found for each of BSE BANEX and BSE Smallcap, the break dates are found to be before demonetization.

Model A				
<i>Variable</i>	Coefficient	Standard error	<i>t</i> -statistic	<i>p</i> -value
<i>Constant</i>	10.073	0.003	3051.431	0.000
<i>Linear trend</i>	0.001	0.00003	38.118	0.000
<i>D</i> <sub>1</sub>	-0.102	0.005	-19.023	0.000
Model B				
<i>Constant</i>	10.103	0.003058	3304.215	0.000
<i>Linear trend</i>	0.001	0.00002	40.931	0.000
<i>D</i> <sub>2</sub>	-0.066	0.004	-15.822	0.000
Model C				
<i>Constant</i>	10.104	0.005	2070.105	0.000
<i>Linear trend</i>	0.0006	0.00004	16.760	0.000
<i>D</i> <sub>3</sub>	-0.00005	0.0001	-0.575	0.566
Model D				
<i>Constant</i>	10.082	0.003	2991.784	0.000
<i>Linear trend</i>	0.001	0.00003	33.038	0.000
<i>D</i> <sub>1</sub>	-0.115	0.005390	-21.408	0.000
<i>D</i> <sub>3</sub>	0.000416	0.0001	6.930	0.000

**Table 3.7A** Estimated Trend Models for BSE SENSEX based on Full Sample

Model A				
<i>Variable</i>	Coefficient	Standard error	<i>t</i> -statistic	<i>p</i> -value
<i>Constant</i>	9.710	0.005	2024.678	0.000
<i>Linear trend</i>	0.002	0.00004	45.470	0.000
<i>D</i> <sub>1</sub>	-0.143	0.008	-18.410	0.000
Model B				
<i>Constant</i>	9.753	0.005	2035.168	0.000
<i>Linear trend</i>	0.001	0.00003	48.656	0.000
<i>D</i> <sub>2</sub>	-0.081	0.007	-12.228	0.000
Model C				
<i>Constant</i>	9.746	0.007	1418.694	0.000
<i>Linear trend</i>	0.001	0.0005	25.163	0.000
<i>D</i> <sub>3</sub>	-0.0003	0.0001	-2.827	0.005
Model D				
<i>Constant</i>	9.716	0.005	1898.499	0.000
<i>Linear trend</i>	0.002	0.00004	39.920	0.000
<i>D</i> <sub>1</sub>	-0.153	0.008248	-18.540	0.000
<i>D</i> <sub>3</sub>	0.0003	0.0001	3.312	0.001

**Table 3.7B** Estimated Trend Models for BSE BANEX based on Full Sample.

Model A				
<i>Variable</i>	Coefficient	Standard error	<i>t</i> -statistic	<i>p</i> -value
<i>Constant</i>	9.715	0.017	555.414	0.000
<i>Linear trend</i>	0.001	0.0001	10.485	0.000
<i>D</i> <sub>1</sub>	-0.161	0.028	-5.677	0.000
Model B				
<i>Constant</i>	9.767	0.016	630.020	0.000
<i>Linear trend</i>	0.001	0.0001	10.088	0.000
<i>D</i> <sub>2</sub>	-0.031	0.021	-1.45	0.149
Model C				
<i>Constant</i>	9.712	0.018	534.401	0.000
<i>Linear trend</i>	0.001	0.0001	10.409	0.000
<i>D</i> <sub>3</sub>	-0.002	0.0003	-5.369	0.000
Model D				
<i>Constant</i>	9.689	0.0187	518.025	0.000
<i>Linear trend</i>	0.002	0.0002	11.160	0.000
<i>D</i> <sub>1</sub>	-0.122	0.030	-4.067	0.000
<i>D</i> <sub>3</sub>	-0.001	0.0003	-3.645	0.001

**Table 3.7C** Estimated Trend Models for BSE Auto Index based on Full Sample

Model A				
Variable	Coefficient	Standard error	t-statistic	p-value
Constant	6.979	0.018	395.202	0.000
Linear trend	0.002	0.0001	17.167	0.000
D <sub>1</sub>	-0.311	0.029	10.797	0.000
Model B				
Constant	7.070	0.015	481.870	0.000
Linear trend	0.001	0.0001	18.521	0.000
D <sub>2</sub>	-0.232	0.020	-11.780	0.000
Model C				
Constant	7.125	0.021	342.966	0.000
Linear trend	0.001	0.0002	4.258	0.000
D <sub>3</sub>	0.001	0.0004	3.847	0.000
Model D				
Constant	7.046	0.017	410.991	0.000
Linear trend	0.002	0.0001	12.559	0.000
D <sub>1</sub>	-0.402	0.027	-14.770	0.000
D <sub>3</sub>	-0.003	0.0003	9.708	0.000

**Table 3.7D** Estimated Trend Models for BSE Reality Index based on Full Sample

Model A				
Variable	Coefficient	Standard error	t-statistic	p-value
Constant	9.192	0.005	1723.344	0.000
Linear trend	0.002	0.00004	36.109	0.000
D <sub>1</sub>	0.107	0.0087	-12.392	0.000
Model B				
Constant	9.223	0.004	2175.726	0.000
Linear trend	0.001	0.00002	53.310	0.000
D <sub>2</sub>	-0.087	0.006	-15.022	0.000
Model C				
Constant	9.240	0.006510	1419.256	0.000
Linear trend	0.001	0.0001	20.005	0.000
D <sub>3</sub>	0.0004	0.0001	3.560	0.000
Model D				
Constant	9.213	0.005030	1831.592	0.000
Linear trend	0.001	0.0001	32.223	0.000
D <sub>1</sub>	-0.139	0.008	-17.217	0.000
D <sub>3</sub>	-0.001	0.0001	10.748	0.000

**Table 3.7E** Estimated Trend Models for BSE Smallcap Index based on Full Sample

Now, reporting on the trend models, we note from the estimated coefficients presented in Tables 3.7A through 3.7E that for all the five models, the constant, time trend and demonetization dummy are significant for all the five series with the exception of the coefficient of  $D_3$  in case of Model C for BSE SENSEX and that of  $D_2$  in case of Model B for BSE Auto.

The impact of demonetization was found to remain effective for a period ranging from 2 to 4 months (January to March 2017) for these series. The coefficient associated with  $D_2$  was found to be negative and statistically significant for these series implying that demonetization led to a fall in stock prices of these series. It may be mentioned here that the basic model that we have considered in case of the Bai Perron test includes a constant and a trend term. Further, all the five series are found to have the coefficient associated with intercept dummy,  $D_1$ , to be significantly negative implying a strong effect of demonetization in pulling down the level values of the series. The slope dummy, however, is found to have differing effects across the five series. For example, the slope dummy has no significant impact on trend for BSE SENSEX, but has strong negative effect for BSE Auto and BSE BANKEX as well as strong positive effect for BSE Realty and BSE Small cap.

### 3.2. Stationary Models for Pre- and post- Demonetization Periods

Apart from the analyses as discussed in the preceding section, an exercise was undertaken to examine if the univariate models for returns on these indices have changed after this important monetary policy decision. The purpose was to find if this very rare major decision which acted like a huge shock to the Indian economy since it was announced all of a sudden without even any inkling at any level by any source, and the impact of which was hitherto unknown and the nation was taken completely off-guard, had any effect in the returns models. This exercise of finding the models with stationary data based on the assumed date has been done from consideration of intervention analysis where the date of intervention is always assumed to be known. Treating this act of demonetization as a major monetary intervention by the Government of India for achieving some very important goals for the Indian economy, we merely wanted to find the impact of this intervention, if any, on the returns model. Accordingly, we obtained the ‘best’ univariate model for each of five returns series for both the pre- and post-demonetization periods. These estimated models are presented in equations (3.1) through (3.8) below.

*Pre-demonetization period (returns on BSE SENSEX)*

$$\hat{r}_{1t} = 0.006 - 0.162\hat{r}_{1,t-4} \quad (3.1)$$

(0.867) (2.406)\*\*

*Post-demonetization period (returns on BSE SENSEX)*

$$\hat{r}_{1t} = 0.007 - 0.169\hat{r}_{1,t-3} \quad (3.2)$$

(1.233) (2.017)\*\*

*Pre-demonetization period (returns on BSE BANKEX)*

$$\hat{r}_{2t} = 0.001 + 0.157\hat{r}_{2,t-1} - 0.16\hat{r}_{2,t-4} + 0.165\hat{r}_{2,t-9} \quad (3.3)$$

(1.15) (2.305)\*\* (2.345)\*\* (2.465)\*\*

*Post-demonetization period (returns on BSE BANKEX)*

$$\hat{r}_{2t} = 0.001 - 0.140\hat{r}_{1,t-3} \quad (3.4)$$

(1.195) (1.657)\*

*Pre-demonetization period (returns on BSE Reality)*

$$\hat{r}_{4t} = 0.001 + 0.135\hat{r}_{4,t-1} - 0.144\hat{r}_{4,t-3} - 0.233\hat{r}_{4,t-4} + 0.144\hat{r}_{4,t-5} \quad (3.5)$$

(0.101) (1.769)\* (1.928)\* (3.059)\*\*\* (1.841)\*

$$+ 0.125\hat{r}_{4,t-7} - 0.137\hat{r}_{4,t-8} + 0.218\hat{r}_{4,t-9}$$

(1.692)\* (1.828)\* (2.879)\*\*\*

*Post-demonetization period (returns on BSE Reality)*

$$\hat{r}_{4t} = 0.003 + 0.193\hat{r}_{4,t-3} - 0.133\hat{r}_{4,t-9} \quad (3.6)$$

(1.634) (2.537)\*\* (1.748)\*

*Pre-demonetization period (returns on BSE Smallcap)*

$$\hat{r}_{5t} = 0.003 - 0.133\hat{r}_{5,t-1} \quad (3.7)$$

(0.382) (1.910)\*

Post-demonetization period (returns on BSE Smallcap)

$$\hat{r}_{5t} = 0.001 + 0.171\hat{r}_{5,t-3} \quad (3.8)$$

(0.171) (2.058)\*\*

Note: No lag value has been found to be significant for returns on BSE Auto Index.

(\*, \*\* and \*\*\* indicate significance at 10%, 5% and 1% levels, respectively.)

It is clear from these equations that in case of four returns series viz., BSE SENSEX, BSE BANKEX, BSE Reality, and BSE Smallcap the lagged structures of the models on returns have changed in the pre- and post-demonetization periods which suggest that the underlying autocorrelation process for each of returns series has changed in these two periods. The ‘best’ models in these two sub-periods were obtained as special AR(4) and AR(3) models in the two sub-periods, respectively, in case of BSE SENSEX returns. It is observed that for returns on BSE SENSEX, the 4<sup>th</sup> lag is significant in pre-demonetization period whereas in the post-demonetization period it is the 3<sup>rd</sup> lag. For BSE BANKEX, lags 1, 4 and 9 are significant at 5% level of significance in the pre-demonetization period, but in the post-demonetization period only the 3<sup>rd</sup> lag is significant, that too at 10% level only. Similarly, the difference in the stationary models in the two sub-periods are found for the returns on BSE Reality and BSE Smallcap indices, as indicated in equations (3.5) through (3.8). Finally, no lagged term has been found to be significant for the periods for returns on BSE Auto, and hence the conclusion is that this series is white noise in both the sub-periods, implying that the change in values of BSE auto index is completely random, and hence in that sense there is no effect of demonetization in case of this returns series.

### 3.3. Change in the Status on Stationarity/Nonstationarity due to Demonetization

Lastly, we now discuss the impact of demonetization in terms of change in the status of the index series in respect of a basic characteristic of any time series viz., stationarity. To that end, we have performed the ADF unit root tests on all the five time series at level values for both the pre-and post- demonetization periods to find whether consequent upon demonetization the status of the series in terms of stationarity/nonstationarity has changed or not. These results are presented in Table 3.8.

Index	Pre-demonetization period		Post-demonetization period	
	ADF test statistic	p-value	ADF test statistic	p-value
BSE SENSEX	-3.551	0.037	-4.241	0.005
BSE BANKEX	-3.475	0.045	-3.383	0.058
BSE Auto	-4.592	0.001	-4.316	0.004
BSE Reality	-2.426	0.365	-3.031	0.12
BSE Smallcap	-4.128	0.006	-4.770	0.000

**Table 3.8** Results of Unit Root Test at Level Values for Pre-and Post- Demonetization Periods

Notes: p-values are taken from Mackinnon (1996); constant and deterministic trend terms have been found to be significant in each series.

It is observed from this table that BSE SENSEX series has a unit root at 1% level of significance and hence it is nonstationary in the pre-demonetized period. However, we also note that the null of ‘unit root’ is rejected at 1% level of significance in the post-demonetization period. But, in the post-demonetization period, the estimate of the coefficient attached to the linear trend term in the ADF estimating equation is found to be 0.0002 which is significant at 1% level of significance. Hence, the series is nonstationary in the sense of being trend stationary only in the



post-demonetization period. Thus, it may be concluded that even after demonetization, the status of this series has remained nonstationary although of trend stationary kind only. Almost similar is the case with BSE Auto and BSE Smallcap. For these two indices, unit root is not significant in both the sub-periods, but the deterministic linear trend has been found to be significant at 1% level of significance. Thus, it may be concluded that there has been no change in this basic property of time series for these two indices even after demonetization with both the series being trend stationary process in both the sub-periods. Insofar as BSE BANEX and BSE Reality are concerned, it is clear from Table 3.8 that both are nonstationary in both the sub-periods. The ADF test statistic values for these two series are -3.475 and -2.426, respectively, in the pre-demonetization period, which are insignificant at 1% level.

Thus both have unit roots in the pre-demonetized period. Same is the conclusion in the post-demonetized period as well for both the series. Based on these findings, we can conclude that insofar as this fundamental property of time series is concerned, demonetization as unanticipated monetary shock has made some impact only in case of one out of the five series considered *viz.*, BSE SENSEX, in the sense that this series has become trend stationary only after demonetization.

#### 4. CONCLUSIONS

This paper has carried out a preliminary time series analysis using daily-level data covering the period 1 January 2016 to 31 May 2017, to study the impacts, if any, of implementation of the important monetary policy of demonetization in India on BSE SENSEX which is the most well-known overall stock index of India and four major financial sectoral indices *viz.*, BSE BANKEX, BSE Auto, BSE Reality, and BSE Smallcap. The findings show that, based on full sample, all the series are integrated of order one implying that all are nonstationary at level values but stationary at their first difference values. The test for existence of multiple structural break points shows that while there is more than one break in all the five series at their level values, there is only one structural break after demonetization.

For all the index series, the constant term, linear trend term and demonetization dummy variable in the five trend models considered have been found to be significant with the sole exception of dummy variable being insignificant in Model C for BSE SENSEX and Model B for BSE Auto. This, of course, shows the impact of demonetization on the trend behavior of all the series. However, no break has been found in the first difference values *i.e.*, for returns, on BSE SENSEX, BSE Auto and BSE Reality. As regards the remaining two returns series *i.e.*, returns on BSE BANEX and BSE Smallcap, one break has been found to be significant for each at 5% level of significance, but both the estimated break dates are before demonetization. Further, some changes have also been observed in the stationary models for the two sub-periods of pre- and post-demonetization -- for all but BSE Auto index. Finally, except for BSE SENSEX series, no change in the status of stationarity/nonstationarity in the pre- and post- demonetization periods has been found in the other series. Even for BSE SENSEX, the series has remained nonstationary in both the pre- and post-demonetization periods, the only difference being that it has become trend stationary only in the post-demonetization period. Overall, it may thus be concluded that except for occurrence of one structural break after demonetization which is also captured in the estimated trend models for all the five series, and some changes in returns model after demonetization for some series, no other major impact of this very important financial decision of demonetization by the Government of India has been found in any of the five financial sectors considered in this study.

## REFERENCES

- Ahmed, M., G. Haider and A. Zaman (2017). Detecting Structural Change with Heteroskedasticity. *Communications in Statistics-Theory and Methods*, 46 (17), 10446–455.
- Andrews, D.W.K. (1993). Tests for parameter instability and structural change with unknown change point. *Econometrica*, 61, 821–856.
- Bai, J., and P. Perron (1998). Estimating and Testing Linear Models with Multiple Structural Changes. *Econometrica*, 66 (1), 47–78.
- Bai, J., and P. Perron (2001). *Multiple and Structural Change Models: A Simulation Analysis*. Columbia University Press.
- Bai, J., and P. Perron (2003). Computation and Analysis of Multiple Structural Change Model. *Journal of Applied Econometrics*, 18, 1–22.
- Chow, G.C. (1960). Tests of Equality Between Sets of Coefficients in Two Linear Regressions. *Econometrica*, 28 (3), 591–605.
- Dasgupta, D. (2016). Theoretical Analysis of Demonetization. *Economic and Political Weekly*, 51 (51), 67 - 71.
- Ioannidis, C. and A. Kontonikas (2006). Monetary Policy and the stock market: Some International Evidence. *Working Paper*. University of Glasgow, Glasgow.
- Kiyotaki, N. and R. Wright (1989). On Money as a Medium of Exchange. *The Journal of Political Economy*, 87 (4), 927–954.
- Maasoumi, E., A. Zaman and M. Ahmed (2010). Tests for Structural Change, Aggregation, and Homogeneity. *Economic Modeling*, 27 (6), 1382–1391.
- MacKinnon, J.G. (1996). Numerical distribution functions for unit root and cointegration tests. *Journal of applied econometrics*, 11 (6), 601–618.
- Muthulakshmi, E.K. and G. Kalaimani (2016). Impact of Demonetization on Indian Economy. *IOSR Journal of Humanities and Social Sciences*, 50–54.
- Nelson, C.R. and C.I. Plosser (1982). Trends and random walks in macroeconomic time series: Some evidences and implications. *Journal of Monetary Economics*, 10, 139–62.
- Rao, K., S. Mukherjee, S. Kumar, D.P. Sengupta, S. Tandon and S. Nayudu (2016). Demonetization: Impact on the Indian Economy. *NIPFP, Working Paper No. 182*. New Delhi.
- Rogoff, Kenneth S. (1998). Foreign and Underground Demand for Euro Notes: Blessing or Curse? *Economic Policy*, 26, 263–303.

- Rogoff, Kenneth S. (2002). The Surprising Popularity of Paper Currency. *Finance and Development* 39 (1), 56–7.
- Rogoff, Kenneth S. (2014). Costs and benefits to phasing out paper currency. *NBER Macroeconomics Annual Conference*. Harvard University Press.
- Rogoff, Kenneth S. (2016). *The Curse of Cash*. Princeton: Princeton University Press.