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DISCUSSION PAPER SERIES

IZA DP No. 14358

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ABSTRACT

Literacy and Information*

Information campaigns aimed at empowering the poor often fall short of meeting their desired aims. We study literacy's role in determining their efficacy. First, exploiting an RD design, we show that receipt of information increased household rice receipts by 30 percentage points. Second, we show that approximately half of the effect is driven by household head literacy. Leveraging novel data on the locations and timings of school openings in the 1970s INPRES school building program, we document that household heads' literacy gained during childhood was pivotal for their households subsequently receiving their full entitlement of rice during adulthood.

JEL Classification: D04, D73, I21, I28, I32, I38, J24, O12

Keywords: poverty, targeting, information, literacy, dynamic complementarity

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1 Introduction

Knowledge is power. Information is liberating. Education is the premise of progress, in every society, in every family

Kofi Annan

7th Secretary-General of the UN

Social welfare programs constitute potentially crucial interventions to provide minimum levels of income and standards of living to poor households and vulnerable individuals. These programs are plagued by exclusion errors however (Ravallion, 2007), wherein many intended beneficiaries fail to receive their full program entitlements. Since beneficiary households may not have full information regarding their program entitlements, incomplete information has been argued to represent key constraints on households receiving their full program entitlements (Banerjee et al., 2018). Concurrently, many poor households globally lack access to basic education, which has been shown to be pivotal in poor household's social and economic outcomes (Duflo, 2001). Greater dissemination of information has therefore been championed as a means to empower the poor, while literacy, in particular, has been singled out as a pre-requisite for information campaigns to be successful (World Bank, 2004; Ravallion et al., 2013).

Testing the role of literacy in fostering successful information campaigns is challenging, since education and information are both endogenous to any social welfare program outcome. Indeed individuals typically acquire their schooling during *childhood*, whereas the effects of information campaigns are likely operate through receipt of information by household heads during *adulthood*. Causally identifying the role of literacy in fostering successful information campaigns therefore necessarily involves linking two causal policy evaluations over the lifetimes of the *same* individuals so as to examine this dynamic complementarity.

To address this challenge, we implement a two-part intertemporal policy evaluation in the context of Indonesia. Our analysis during household heads' adulthoods links the provision of information with household receipts from a social welfare program in the context of *Raskin* rice (see [Banerjee et al., 2018](#)). We subsequently show that the efficacy of the information treatment is in large part driven by household heads' literacy. Our analysis during household heads' childhoods, causally links one of the largest school building programs to household head literacy in the context of the *INPRES* school building program a la [Duflo \(2001\)](#). Ultimately therefore, we are able to provide a causal chain of results, linking: the information treatment to receipts of *Raskin* rice, the information treatment and household head literacy (to show the heterogeneity in households' receipt of *Raskin* rice) and the building of *INPRES* schools to household head literacy in childhood.

The setting for our first evaluation, is the 2014 campaign launched by the Government of Indonesia (GoI) that constituted one of the largest information interventions in the history of poverty reduction programs (see: [Banerjee et al. \(2018\)](#)). Treated households were delivered an information package detailing their social welfare program entitlements (See Figure B1 in the Appendix). Our data detail individual household's PMT scores that the GoI used to identify households' eligibility. The availability of all 482 official Proxy Mean Test (PMT) district eligibility thresholds in turn motivates a RD design, the results of which we corroborate with both parametric and semiparametric methods. Our results show that treated households receive 30 percentage points more *Raskin* rice.¹

We subsequently evaluate the impact of the *SD INPRES* program on household heads' literacy acquired during their childhood. The *SD INPRES* program constitutes one of the largest school construction programs in the world to date ([World Bank, 1990](#)). During the period 1971-1974, the Indonesian government built approximately 61,000 new schools, one new school per 500 children ages 5 to 14 ([Duflo, 2001, 2004](#)), which is equivalent to a doubling of the stock of primary schools nationally (please refer to [Figure 3](#)).

We leverage novel administrative data from Indonesia's Ministry of Education on

¹In doing so, we provide external validity to [Banerjee et al. \(2018\)](#) in the sense that we use nationally representative data as opposed to an RCT across six out of 514 municipalities.

school information, namely *DAPODIK* (for *Data Pokok Pendidikan Dasar dan Menengah* or Basic Data on Primary and Secondary Education), to identify the village primary schools that each of our household heads (receiving our information treatment) were exposed to during their childhood. We subsequently leverage the time and spatial variation in our school building data to construct instruments for childhood exposure to the *SD INPRES* program, having accounted for migration since birth, to causally link the *SD INPRES* program to literacy in adulthood. The results suggest that having graduated from a primary school during childhood increases the probability of the household heads being literate by 49.4 percentage points.

Taken collectively, we provide causal evidence that literacy constitutes a key mechanism that determines the efficacy of information campaigns; as measured in our case by households' receipts of *Raskin* rice. In doing so, we link two seminal papers from across the literature, [Duflo \(2001\)](#) and [Banerjee et al. \(2018\)](#), to demonstrate how treatments can affect one another intertemporally. In contrast to [Banerjee et al. \(2018\)](#) we interact household head literacy with our RD estimates, to causally show that approximately half of the observed effect of information on the receipt of *Raskin* rice is driven by household head literacy. We therefore highlight how policy evaluations may complement one another dynamically for the same individual over their lifetimes along the intensive margin, as opposed to complementing one another when delivered in tandem along the extensive margin (see [Tohari et al., 2019](#)). In doing so, we contribute to the burgeoning literature on dynamic complementarities, see for example: [Cunha and Heckman \(2007\)](#); [Chetty et al. \(2011\)](#); [Aizer and Cunha \(2012\)](#); [Attanasio \(2015\)](#); [Attanasio et al. \(2017\)](#); [Foster and Gehrke \(2017\)](#); [Johnson and Jackson \(2019\)](#).

Our research speaks directly to the literature that examines the role of information campaigns on the uptake of social welfare programs, which to date has delivered mixed results ([Olken, 2007](#); [Reinikka and Svensson, 2004](#); [Pandey et al., 2009](#); [Banerjee et al., 2010](#); [Pradhan et al., 2014](#); [Lieberman et al., 2014](#)). In our context of anti-poverty programs, [Ravallion et al. \(2013\)](#) finds that benefits are *negatively* associated with education. [Banerjee et al. \(2018\)](#), in evaluating the same information treatment as we do in this paper

through implementing field experiments, rather use household head literacy to balance their treatment and control groups. It remains unclear however, why some information-based interventions succeed, while others do not. One plausible explanation pertains to the extent to which information is understood by eligible households (Fox, 2007). Indeed, the World Bank (2004) claims that there is a strong relationship between the success of information treatments and levels of public literacy, especially among poor people. Our results provide causal evidence in favor of this explanation.

We also contribute to the rich and longstanding literature that examines the short, medium and long-term impacts of the *SD INPRES* program on: labour force productivity and economic growth (Duflo, 2001, 2004), improvements in local governance and public good provision (Martinez-Bravo, 2017), health outcomes (Breierova and Duflo, 2004; Somanathan, 2008) and intergenerational effects (Akresh et al., 2018; Mazumder et al., 2019). While existing studies examine the *SD INPRES* program using aggregate schooling data however, (Duflo, 2001, 2004), our analysis estimates the effect of the school construction at the village level which are commonly used as catchment areas for primary schools in Indonesia (see for example Bazzi et al. (2020)). Whereas Martinez-Bravo (2017) use the school data from PODES which are also disaggregated at the village level, those data are silent as to when schools were constructed. Indeed, many schools were built prior to the *SD INPRES* period, which if left unaccounted for may bias estimates. We leverage the fact that between 1971 and 1990 the *SD INPRES* program reduced illiteracy across Indonesia from 39.1 % to 15.8% Duflo (2001, 2004), a fact that we exploit once household heads have matured into adulthood. Taken collectively, we interpret our results as confirmation of the role of past (mass) education policy in improving the effectiveness of contemporaneous anti-poverty programs.

2 Institutional Background

2.1 Pre-Information Campaign Performance of Raskin

Since 1997, the Government of Indonesia has implemented several strategies and programs to alleviate poverty (see [Tohari et al., 2019](#)). These programs are clustered around their targeted beneficiaries. Programs targeted at individuals (e.g. *Jamkesmas*)² and households (e.g. *Raskin*, *BLSM*, and *PKH*)³ comprise the first cluster. Community targeted programs (e.g. *PNPM Mandiri*)⁴ fall under the second. The third cluster includes programs targeted at micro and small enterprises (e.g. *Kredit Usaha Rakyat – KUR*).⁵

Previous research has identified several program deficiencies, which for *Raskin* include: (1) Rice not reaching eligible households, i.e. exclusion errors during the delivery process⁶ and (2) Evidence of frequent *Raskin* purchases by poor and non-poor households alike ([Olken, 2005](#); [Banerjee et al., 2018](#)),⁷ a fact that underpins our identification strategy; and (3), Local governments failing to judiciously allocate the *Raskin* budget thereby leading poor households to pay higher prices for rice in addition to delays in rice distribution

²*Jamkesmas* is health insurance for the poor (previously known as *Asuransi Kesehatan untuk Keluarga Miskin*, or *Askeskin*, later renamed *Jamkesmas*). In 2014, *Jamkesmas* covered some 24.7 million households or 96.4 million people.

³*PKH* is a Conditional Cash Transfer program managed by the Indonesian Ministry of Social Affairs that targets the bottom 5% of the population. PKH beneficiaries receive direct cash transfers ranging from IDR. 600,000 to IDR. 2.2 million or (about USD\$67–\$250) depending on their family composition, school attendance, pre-/postnatal check-ups and vaccination completions.

⁴*PNPM Mandiri* (for *Program Nasional Pemberdayaan Masyarakat Mandiri* or the National Program for Community Empowerment) is Indonesia’s largest community-driven development program to help alleviate poverty through empowering local communities. There are several components of the *PNPM Mandiri*, two of which are PNPM Rural, that began in 1998 as *Kecamatan* Development Program (KDP) and PNPM Urban, which begun in 1999 as the Urban Poverty Program (UPP). Interested readers are referred to [TNP2K \(2015b\)](#).

⁵*KUR* (for *Kredit Usaha Rakyat* or credit for micro and small enterprises) are credit/working capital and/or investment financing schemes for enterprises that are unable to meet certain banking requirements. The amount of credit provided to each enterprise is less than IDR. 5 million (about \$500).

⁶Existing administrative records are unable to indicate the point at which the “missing” rice exits the delivery chain since no single authority is responsible from the point of *Raskin* rice procurement to household purchase ([World Bank, 2012](#)).

⁷The amount of *Raskin* rice purchased by a household is roughly constant across the entire consumption distribution, meaning non-poor households buy as much *Raskin* as poor, near-poor, or vulnerable households ([World Bank, 2012](#)). In 2010, the [World Bank \(2012\)](#) estimates that the average amount of *Raskin* rice bought by poor households was approximately 3.8 kilograms per month.

(Hastuti et al., 2012).

To address these shortcomings, between 2011 and 2014, the GoI made significant changes to the targeting mechanisms and service deliveries of several anti-poverty programs including *Raskin*. Specifically, the Unified Data Base (UDB) was developed to identify the poorest 40% of the population for inclusion in social assistance programs through Proxy Means Testing (PMT) (see Tohari et al., 2019, for detailed discussions of the targeting improvement). Following improvements in targeting, in the third quarter of 2013, the GoI also distributed Social Security Cards (*Kartu Perlindungan Sosial - KPS*) which signify beneficiaries' eligibility for anti-poverty programs, in addition to an information packet, the contents of which constitute our first treatment.

2.2 The Information Intervention

The information treatment that we analyse was delivered directly to households using the postal service. As shown in Figure B.1 of Appendix B, the information provided details on accessing three different social welfare programs to which the KPS card entitled households, in addition to the complaints procedure to follow should households feel unfairly treated. Due to the nature of the information intervention which includes significant text, *a priori* one can expect household head literacy to matter in terms of their understanding the information provided. Our outcome variable however, in line with Banerjee et al. (2018), is solely on the amount of *Raskin* rice purchased by households.⁸

This information campaign (and the rollout of the KPS card) was targeted at the bottom 25 percent of households, equivalent to 15.5 million poor and near-poor Indonesian households, the names and addresses of which were taken from the UDB (see Figure 1). This intervention was delivered directly to targeted households by *PT POS Indonesia*, the State-owned postal company.

⁸Details of the KPS card are provided in Appendix Figure B.2 In an effort to protect the card from fraud, the KPS card includes household details including: household head, their spouse and address as well as barcodes representing the family card number.

2.3 *Raskin* Delivery Mechanism

The *Raskin* program aims to reduce household expenditure on food, particularly on rice, the staple food in Indonesia. In 2013 and 2014, the program covered around 15.5 million of the poorest Indonesian households. According to the 2014 *Raskin* Guidelines,⁹ the implementation of the program has not changed since its inception. Figure B.3 in the Appendix B shows the delivery mechanism for the *Raskin* program. Since 2011, several agents have been involved in the procurement and delivery of *Raskin* rice. They include: (i) the Coordinating Minister of Social Affairs (for *Kementerian Koordinator Bidang Kesejahteraan Rakyat* or Coordinating Minister of Social Affairs), later called *Kemmenko PMK* (for *Menteri Koordinator Bidang Pembangunan Manusia and Kebudayaan* or Coordinating Minister of Human Resources and Culture), and the Vice President’s National Team for the Acceleration of the Poverty Reduction (TNP2K), which together determine yearly allocation and price of rice,¹⁰ (ii) the *Bulog* (the National Logistics Agency) responsible for procuring rice from producers and delivering the rice to over 50,000 distribution points across Indonesia. *Raskin* beneficiaries are expected to make monthly *Raskin* purchases from these distribution centres¹¹ and (iii) the District government that is responsible for the logistics of transporting *Raskin* rice to recipient households.

We measure the effectiveness of the information intervention using the average amount of *Raskin* rice bought by beneficiary households over a three-month period. Summary statistics of our outcome as well as the characteristics of *Raskin* beneficiaries are presented in Appendix Table A.1 On average, households in receipt of the information treatment bought around six kilograms of *Raskin* rice, which is less than half the intended allocated

⁹Kemenkokesra. (2014). “*Pedoman Umum Raskin 2014*” (General Guideline: Rice Subsidy for Poor People 2014). Jakarta: Kemenkokesra.

¹⁰According to the general guidelines of *Raskin* 2014, the total number and the list of *Raskin* beneficiaries were obtained from the Unified Database of TNP2K. In terms of benefit, each targeted household should receive 15 kg/month per month of rice. The price of *Raskin* rice is IDR 1600 /kg at the Sharing Point (*Titik Bagi*).

¹¹The distribution centres (or *Titik Distribusi*) of *Raskin* are mostly located in village offices or other places that are decided upon between Local Government and Bulog. The local government and village administrative apparatuses are then responsible for notifying eligible beneficiaries and arranging the transport of rice from distribution points to households (*Titik Bagi* or sharing points).

benefit; although this is significantly higher than the purchase of *Raskin* rice by those households in 2011, which was only 3.7 kilograms.

Although, as shown in [Figure 1](#), the bottom 25% of the population of Indonesia were *de jure* eligible for *Raskin* rice purchases in addition to the information treatment, *de facto*, households on both sides of this threshold were also able to purchase *Raskin* rice, which resulted from the actions of village leaders. To substantiate this claim, [Figure 2](#) presents both the proportion of households purchasing *Raskin* rice, as well as the average amount of *Raskin* rice purchased by consumption decile in 2011. Crucially, from the perspective of our identification strategy, which focuses on those just above and below the threshold of the poorest 25% of households, as shown by the dashed (red) line, we do not observe any significant differences in the amount of *Raskin* rice purchased between these two groups. In other words, while there are no discernible differences in *Raskin* purchases above and below our threshold, those below our 25% threshold were targeted by the information campaign, while those above the threshold were not. It is this discontinuity that we exploit in the first part of our analysis.

3 Data, PMT Score and Eligibility

To evaluate the effect of the information campaign on the amount of *Raskin* rice received, we combine several nationally representative surveys with rich administrative data obtained from the GoI.

3.1 The *SUSENAS* Survey

The National Socioeconomic Survey (SUSENAS) is an annual cross-sectional, nationally representative dataset, initiated in 1963-1964 and fielded annually else biannually since then. In 2011, the Central Bureau of Statistics of Indonesia (BPS) changed the survey frequency to quarterly, and for each quarter, the SUSENAS covers some 300,000 individuals and 75,000 households. In this paper, we utilize data from the 2014 wave of the SUSENAS survey to: (i) generate variables that are required to estimate the PMT Score

for each household using the official PMT coefficients (ii) obtain control variables that are not included in the PMT score estimation and (iii) construct poverty indicators as outcome variables.

3.2 Social Protection Survey (SPS)

The second dataset used in our analysis is the 2014 Social Protection Survey (SPS), which was conducted jointly by the BPS and TNP2K as a supplement to the SUSENAS. This survey was implemented from the first quarter of 2013 to the first quarter of 2014 and was specifically aimed at examining the performance of poverty targeting following the implementation of the UDB. We exploit data from the first quarter of 2014, the period immediately following our information treatment to construct our treatment and outcome variables.

3.3 Village Census (PODES)

We use the 2011 and 2014 waves of PODES data, to provide information on all villages/*desa* in Indonesia. The variables produced using this census include the characteristics of the village, some of which were used in estimating the PMT scores.

3.4 DAPODIK

DAPODIK (for *Data Pokok Pendidikan Dasar dan Menengah* or Basic Data on Primary and Secondary Education) is the administrative school-level dataset collected by the Ministry of Education. *DAPODIK* comprises more than 512 thousand primary and secondary schools and provides details of schools': precise location, date of construction, operational and financial assets etc.

Between 1973 and 1983, the Government of Indonesia embarked on a massive school construction program to increase primary school enrolment, namely the *SD INPRES* program (for *Sekolah Dasar Instruksi Presiden* or Presidential Decree Primary School), one of the largest such programs in history. The schools were constructed in provinces with low primary school enrolment corresponding to one new school per 500 children

aged 5 to 14 in 1971 (Duflo, 2001, 2004). The approximately 61,000 new schools constructed through this program, constituting some 44% of the total, were approximately proportional to the number of school-aged children not enrolled in school prior to the program (Martinez-Bravo, 2017). Each new school was allocated three teachers and 120 primary school pupils, aged between 7 and 12 years.¹² Taken together, Figures 3 and 4 highlight the significant variations in primary school construction in Indonesia, across both time and space. First, we assume that the village household heads currently reside in are the same villages as where they went to primary school in when aged six years old.¹³ Subsequently, we exploit the spatial and time variations in the *SD INPRES* school construction program, in order to define a dummy variable equal to one should a particular household head have had access to a primary school when they themselves became of primary school age.

3.5 Merging the datasets

The greatest challenge in merging our datasets is the fact that since 2011, the BPS has not published the village and subdistrict codes for their household survey data. To address this, we proceed as follows:

- i). We merge the Quarter 1 2014 SPS data with the Quarter 1 2014 SUSENAS using the household ID available in both datasets. Overall, 70,336 households from the SPS sample can be identified from the total of 71,051 households in the SUSENAS survey.
- ii). These combined data are then merged with the 2014 pooled SUSENAS to obtain village and sub-district IDs using a ‘bridging code’ privately shared with us by the BPS.¹⁴

¹²Based on the *DAPODIK* Data, 63,518 primary schools were built through the INPRES Program (from 1974 to 1987), representing about 44% of the total primary schools in Indonesia (total=145,782).

¹³In the SUSENAS 2014, the BPS did not collect any information on migration. In 2019 SUSENAS, however, questions about migration were included in SUSENAS Core questionnaire, and there are 20.23 percent of the population who lived in a municipality different from their birthplace municipality.

¹⁴We are grateful to a TNP2K targeting team who provided us with this bridging code.

- iii). We subsequently merge the resulting dataset with selected variables from PODES using a village identifier so as to obtain village-level variables. After merging with the PODES data, we are able to identify 67,118 households including details of their expenditure and social protection as well as village-level information that can be combined with the official PMT coefficients in order to obtain individual household PMT scores, which are discussed in detail below.
- iv). Finally, using village identifiers from PODES, we construct village polygons using an official map obtained from the BPS. We then conduct spatial merges to overlay all these polygons with the geolocations of all schools detailed in the *DAPODIK* dataset. From this process, we know the exact catchment area for each school as presented in [Figure 4](#).

3.6 Estimating the Household’s PMT Score and their eligibility

Measuring household PMT scores is crucial in defining the eligibility status of each household to our information treatment. Estimating the PMT score involves:

1. Selecting the poorest 25% or 15.5 million households from the UDB. The UDB contains information on the bottom 40% of the Indonesian population collected through PPLS11 (*Program Pendataan Perlindungan Sosial 2011*) together with their estimated PMT scores. To estimate the PMT score and rank of each household in the UDB, the GoI used coefficients that are measured using SUSENAS and PODES 2011. These coefficients are unique to the 482 districts from the total of Indonesia’s 497 districts in 2011.¹⁵ The PMT score for each household is then measured using each household’s observable information, which in turn is plugged into the corresponding district coefficient and subsequently ranked. Using household’s PMT scores and ranks, the government then selects a list of intended beneficiary households.

¹⁵For other 15 districts, the GoI implement universal targeting system. For these specific areas, such as several districts which have high incidence of poverty, the GoI selects intended beneficiaries using a ‘negative lists’ method, which means all households are eligible for poverty programs, except those that contain a public servant, local leaders, high ranking military officials etc.

2. Using these official PMT coefficients, this study recovers households' PMT scores in 2014: (1) using data from SPS, SUSENAS and PODES in 2014, to construct variables that are comparable to those variables used in PPLS11 (2) following the same steps as conducted by the GoI in which the 2014 variables are plugged into the official PMT coefficients and (3) ranking each household based on their PMT score. As our study uses nationally representative data, the household rank represents their rank relative to the total population. Each household's eligibility for social welfare programs depends upon whether their PMT score lies above or below their district's cut-off. The cut-off for each district is measured using the official quota used by the GoI to select the list of KPS program beneficiaries that are unique to each district.

We plot the result of this process, the estimates of our PMT score against the probability of receiving the information treatment using a nonparametric [Fan \(1992\)](#) regression estimation in [Figure B.4](#) in the Appendix. The Figure confirms an overall negative relationship between the PMT Score and the likelihood of receiving information intervention, since higher PMT implies lower eligibility for the program.

4 Estimation Strategy

Household's eligibility for the information treatment is based on their household PMT score relative to their district's cut-off. We investigate the impact of receiving the information treatment on household's receipts of Raskin rice. Let $pmt_{i,d}$ be the PMT score for each household and \overline{pmt}_d be the PMT cutoff for each district. Then, I, defines the eligibility of each household to receive the information intervention, ([TNP2K, 2015a](#)):

$$Pr(I = 1) = \begin{cases} 1 & \text{if } pmt_{i,d} \leq \overline{pmt}_d \\ 0 & \text{if } pmt_{i,d} > \overline{pmt}_d \end{cases} \quad (1)$$

For each eligible household, we can define their potential outcome, B , with (B_1) if they received the treatment and (B_0) otherwise. Following [Rubin \(1974\)](#), the difference between the average benefit of recipient households relative to non-treated households

becomes:

$$\mathbb{E}(B | I = 1) - \mathbb{E}(B | I = 0) = \underbrace{\mathbb{E}(B_1 - B_0 | I = 1)}_{\theta} + \underbrace{\mathbb{E}(B_0 | I = 1) - \mathbb{E}(B_0 | I = 0)}_b \quad (2)$$

Our estimate of interest is the average treatment-on-the-treated, i.e, the effect of receiving the information treatment, θ , for a subgroup of compliers. The main threat to identification is the prospect of omitted variable bias, b ; unobserved determinants that are potentially correlated with the probability of receiving the information as well as with the level of benefits received.

4.1 The Impact of Information on the Benefit Received

First we implement a regression discontinuity methodology by exploiting the discontinuity in the eligibility of information provision around our 25% threshold, as in [Equation \(1\)](#). The baseline analysis estimates the following regression model within a narrow window around the district cut-off:

$$\begin{aligned} Out_{i,d} = & \beta_0 + \beta_1 Information_{i,d} + \beta_2(\overline{pmt}_d - pmt_{i,d}) + \beta_3 Information_{i,d} * \\ & f(\overline{pmt}_d - pmt_{i,d}) + \beta_4(1 - Information_{i,d}) * f(\overline{pmt}_d - pmt_{i,d}) + \varepsilon_{i,d} \end{aligned} \quad (3)$$

Where $Out_{i,d}$ is the average amount of Raskin rice bought per month in kilograms by household i in district d . $Information_{i,d}$ is an indicator variable equal to 1 if household receive information treatment, and $\overline{pmt}_d - pmt_{i,d}$ is the distance between households' PMT scores and their respective district cut-offs. We examine alternative functional forms, $f(\cdot)$, of the RD polynomial for robustness. The baseline specifications follow [Gelman and Imbens \(2019\)](#) and implement RD polynomials of differing orders. First however, we show that the information treatment had a significant increase in the amount of rice bought from the *Raskin* program in 2014, for those households located within the RD envelope, see [Figure 5](#).

[Table 1](#) presents our baseline results from estimating [Equation \(3\)](#). To select the optimal bandwidth, we follow the criteria proposed by [Imbens and Kalyanaraman \(2012\)](#)

henceforth, IK2012, in the first three Columns and [Calonico et al. \(2014\)](#), henceforth, CCT2014, in Columns four to six. The polynomial order, the size of the bandwidth and the observations inside the bandwidth are presented in [Table 1](#).

The 2SLS coefficients using nonparametric estimates without adjusting for covariates, in Columns (1) and (4) in Panel A of [Table 1](#), show that in general, receiving information increases the amount of *Raskin* rice purchased by around 30.6 percentage points according to IK2012, and 39.2 percentage points according to CCT2014. We also test whether the treatments differed between Java and Non-Java, by splitting the sample. Java is the most populous island in Indonesia and previous studies (e.g. [Ravallion and Dearden, 1988](#)) have shown that Java tends to be more egalitarian whereby benefits are more often shared. Given this, the distribution of the benefits received from poverty programs could differ between Java and other areas of the country. The results using linear order polynomials in Panel A of [Table 1](#) are presented in Columns (2) and (3) based on IK2012, and Columns (5) and (6) results are based on CCT2014. They show that there are indeed significant differences between the impact of information dissemination in Java and the other provinces, even though the effects are not statistically significant using lower order of polynomials. When we implement cubic order polynomials however, the results in both Java and Non-Java become statistically significant. Under this specification, the effect of information on the Java sub sample is about 61.1 percentage points higher and statistically significant, while the effects in the Non-Java provinces are about 32.8 percentage points under IK bandwidths. Our estimates using higher order polynomials, those excepting cubic order polynomials under the CCT bandwidth selection, likely generate higher estimates because higher order polynomials assign far greater weights to observations further away from the discontinuity [Gelman and Imbens \(2019\)](#).

We also include pre-intervention covariates related to village and head of village characteristics following [Frölich \(2007\)](#) Frolich (2007) and [Calonico et al. \(2019\)](#).¹⁶ [Imbens and Kalyanaraman \(2012\)](#) however, note that the inclusion of additional covariates should not significantly affect such analyses. The results presented in panel B of [Table 1](#), shows

¹⁶Pre-intervention covariates related to village and head of village are derived from 2011 PODES data.

that in general the inclusion of covariates produces slightly lower estimates. For example, using linear order polynomials and the IK bandwidth selection, the covariates-adjusted estimates of providing information on *Raskin* are about 25.9 percentage points higher, while under non-adjusted covariates estimation it is about 30.9 percentage points.

Interestingly the covariates-adjusted RDD estimation under IK2012 bandwidth selection and linear order polynomial produces the closest estimate when compared to the results of [Banerjee et al. \(2018\)](#). Those authors find that the same information treatment increases the amount of *Raskin* rice purchased by approximately 26% when compared to the control group. Since we exploit nationally representative data, as opposed to basing our estimation on an RCT in 550 villages across six districts, we argue that our research provides external validity of [Banerjee et al.](#)'s results.

4.2 Robustness Checks and Extensions

4.2.1 Sensitivity Tests

First, we choose a range of placebo cut-offs to ensure that the discontinuity of the outcome of interest only occurs at the true cut-off. [Table 2](#) summarizes the estimate of the effect of information for selected cut-offs ranging from -0.1 to 0.1 in increments of 0.05. [Figure 6](#) plots the estimates. The cut-off at 0 is included as a benchmark. As expected, with the exception of 0 i.e. the true cut-off, there is no effect of the information treatment at any of the placebo cut-offs. In terms of magnitude, the effect of information is smaller compared to the true effects at all other cut-offs. This implies that the outcome of interest does not jump discontinuously at any other cut-off other than at zero.

In choosing a bandwidth, it is critical to consider an optimal balance between estimation precision and estimation bias ([Lee and Lemieux, 2010](#)). Larger bandwidths, on the one hand, yield more precise estimates since more observations can be relied upon in estimation (i.e. greater efficiency). On the other hand, when a larger bandwidth is used, the resulting estimates are less likely to be accurate as increasingly more observations are considered that are located further from the threshold (i.e. greater bias). [Figure 7](#) plots the estimated 2SLS coefficients of the effect of information and the associated confidence

intervals for different bandwidth selections or window lengths using IK2012. The area within the vertical dashed lines represents the location of the true optimal bandwidths that are selected based on both IK2012 and CCT2014. Evidentially, as the bandwidth increases, the bias of the estimator increases as its variance decreases. Therefore, it is natural that the larger the bandwidth, the smaller the confidence intervals, but due to bias, the point estimates are also displaced.

4.2.2 Comparing RD, LATE and LARF

The results from the local kernel regression results confirm that receiving information significantly increases the benefits received from the *Raskin* program. Below we examine whether these effects are also consistent if they are estimated following Angrist et al. (1996) parametric estimate and Abadie (2003) semiparametric approach.¹⁷ Our parametric approach, the estimation of the LATE, implements an instrumental variable technique with eligibility status of the household used as our instrument for treatment. Our semiparametric approach as detailed in Abadie (2003), instead proposes to use a Local Average Response Function (LARF) that allows one to compare the characteristics of treated and non-treated individuals within the compliers' subset, in the absence of knowledge as to who is and is not a complier. The estimation of the LARF is conducted in two steps which are: (1) to measure weights, w , by estimating parametrically (or non-parametrically) $p(Z = 1|X)$ and (2) estimating the effects using Weighted Least Square (WLS) with weights equal to w .

With regards national level effects, Columns (2) and (5) in Table 3, show the results from both our parametric and semiparametric estimators, which are slightly different and statistically significant. The magnitude of the effects and their signs show that the provision of information increases the benefits received from the *Raskin* program by about 37.1 percentage points in parametric and 48.5 in semiparametric estimations, respectively. The result of parametric estimation is in the range of the estimated effects from our

¹⁷Lee and Lemieux (2010) note a number of alternative estimation strategies and suggest that no single method be relied upon. Our parametric and semiparametric estimations are therefore included to complement our non-parametric approach.

nonparametric approach in [Table 1](#), while the result of semiparametric estimation is slightly higher in all nonparametric alternative estimations.

The difference in the effects of the information treatment between Java and Non-Java is noteworthy. In general, our parametric and semiparametric estimates produce consistent results with the nonparametric estimation in which the effect of information on social benefits away from Java is lower than in Java itself and all the results are statistically significant. In terms of the magnitude however, using our parametric results in Columns (3) and (4) of [Table 3](#), we observe that the provision of information increases the benefits received from the *Raskin* program by about 42.6 percentage points in Java households and by 36.8 percentage points in Non-Java households, respectively. Moreover, our semiparametric results for Java and Non-Java households, produce the same results with small difference between Java and Non-Java compared to our parametric results.

Finally, it is important to note that the OLS estimate in Column (1) of [Table 3](#) is downwardly biased. According to the OLS result, the increase in the benefits received from *Raskin* is about 21.5 percentage points conditional on covariates. The estimated effect of information increases when we instrument this variable with the household's eligibility to receive treatment. Overall therefore, we can conclude that the provision of information to eligible households increases the level of benefits received by between 30-40 percentage points on average.

4.3 Information Intervention and Household Head Literacy

Knowing that the information treatment has a positive impact on the benefit received from *Raskin* program, in this section, we investigate whether household characteristics have contributed to explaining variation in the level of benefits received. As pointed out by [Tohari et al. \(2019\)](#) one of the aims of KPS and information intervention was to increase the complementarities among poverty program in Indonesia. In that paper we show that the implementation of the KPS has increased the complementarities of the programs related to their extensive margin. In this section, we further investigate whether household head literacy contributes to the effectiveness of the information intervention,

in effect providing a formal test of the conclusions of [World Bank \(2004\)](#); hypothesising that literate household heads receive more of their entitlements in comparison with illiterates. This is consistent with the argument that information can only improve public participation and increase benefits if the information is understandable and actionable ([Fox, 2007](#)). We therefore modify [Equation \(3\)](#) and estimate the following regression:

$$\begin{aligned}
Out_{i,d} = & \beta_0 + \beta_1 Information_{i,d} + \beta_2(\overline{pmt}_d - pmt_{i,d}) + \\
& \beta_3 Information_{i,d} * f(\overline{pmt}_d - pmt_{i,d}) + \beta_4 Information_{i,d} * IND_{i,d} + \\
& \beta_5(\overline{pmt}_d - pmt_{i,d}) * IND_{i,d} + \beta_6(Information_{i,d}) * \\
& f(\overline{pmt}_d - pmt_{i,d}) * IND_{i,d} + \varepsilon_{i,d}
\end{aligned} \tag{4}$$

where $IND_{i,d}$ represent indicator variables which equal 1 if households i in the district d have specific characteristics, for example, they receive other programs than *Raskin*, have head or wife who can read, and other characteristics and 0 otherwise.

[Table 4](#) presents the result of whether a household head's literacy explains the heterogeneity in the impact of information intervention on the amount of rice purchased by poor households. Column 1 reproduces the baseline result as presented in the previous section. Column 2 examines whether the effect of information is larger if the household head is literate, which is plausible since literacy is a necessary condition for households to understand the contents of the treatment, through the addition of an interaction term comprising a dummy variable that is equal to one should the household head be literate. The coefficient on the interaction term shows that approximately half of the total effect of the information treatment operates through the channel of household head literacy. We continue by evaluating the comparable effect for household head's wives. The results are reported in column 3 and show that the overall effect of the information treatment is even larger and statistically significant when the wife is literate, although the interaction effect is slightly smaller than the comparable estimate of household head literacy. In summary, literacy is shown to be a critical factor determining the effectiveness of our information treatment.

4.4 Literacy and Past Education Policy

Since household literacy is shown to have a significant impact on the effectiveness of information campaigns, in this section, we examine the effect of past education policy on household literacy. Specifically, we exploit the massive school construction program to increase primary school enrolment, namely *SD INPRES* program (for *Sekolah Dasar Instruksi Presiden* or Presidential Decree Primary School) which was implemented between 1973 and 1983. The schools were specifically constructed in provinces with low primary school enrolment. [Figure 3](#) shows the development of primary school in Indonesia from 1890 to 1990. The figure confirms that during the *INPRES* program, there had been a significant increase of primary school building in Indonesia, equivalent of around 44% of the total stock as of 1990.

We therefore estimate the LATE regression following [Angrist et al. \(1996\)](#) as follows:

$$\begin{aligned} Lit_{i,d} &= \beta_0 + \beta_1 Grad_{i,d} + v_{i,d} \\ Grad_{i,d} &= \alpha_0 + \alpha_1 School_{i,d} + \varepsilon_{i,d} \end{aligned} \tag{5}$$

Where $Lit_{i,d}$ is an indicator variable equal to 1, if individual i in village d is literate, while $Grad_{i,d}$ is an indicator variable equal to 1 if an individual graduated from primary school. Since $Grad$ is an endogenous dummy variable, we instrument that variable with $School$, an indicator variable equal to 1 if a primary school construction had been completed in household (or wife) heads' villages, when they first entered primary education, when he/she was 6 years old.

In order to construct the $School$ variable, we use a novel dataset known as *DAPODIK* (for *Data Pokok Pendidikan* or Education database) which has been collected by the Indonesian Ministry of Education since 2014. This database contains rich information on schools in Indonesia including information: on their exact location transformed into geolocations, when the school was built and first operationalized etc. This estimation strategy is at odds with those operationalised in previous studies that examined the impact of the *SD INPRES* program, which relied upon aggregated data (e.g. [Duflo, 2001, 2004](#)). The additional levels of detail in our data allow us to examine the impact

of school construction on individuals in the catchment areas of specific primary schools. It is common that the primary school catchment in Indonesia is at the village level.

Table 5 reports the results concerning the impact of graduating from an *SD INPRES* primary school on the literacy of household heads (along with their wives). Columns 1 and 4 report the results from our baseline OLS estimates. Columns 3 and 6 rather present the average treatment effects of graduation on literacy rates. Household heads graduating from primary school increased the likelihood of male household heads being literate by some 49.4 percentage points. For their wives, the effect is even larger, since the probability of being literate rose by some 73.3 percentage points. Nevertheless, migration since childhood could bias our results (Duflo, 2004). For the sake of robustness, we therefore estimate Equation 5 separately for migrant and non-migrant households. Table A.2. and Table A.3. report the results. While our overarching results hold, we find that the magnitudes of our estimated effects are somewhat attenuated for household heads and strengthened for the wives of migrant households; while the converse also holds true for non-migrant households.

5 Conclusion

Information campaigns have been proffered as low cost interventions to empower the poor. We contribute to the limited evidence base on their efficacy, evaluating the previously untested hypothesis that literacy is pivotal to their success (World Bank, 2004; Fox, 2007). Such an evaluation is non-trivial, since literacy is typically acquired during childhood, while receipt of information campaigns is most often received during adulthood.

The setting for the first part of our policy evaluation is the 2013 targeted information campaign of the GoI, one of the largest information interventions in history. In evaluating this information campaign, we draw upon several extraordinarily rich sources of administrative data, including GoI PMT coefficients that we use to recover each household's PMT score and which we subsequently compare to that household's district PMT threshold in order to definitively pin down which households *should* have received the treatment. We

implement in an RD framework. We find that the provision of information increased households' receipts of *Raskin* rice by between 30 and 40 percentage points. In doing so, we provide external validity to the results of [Banerjee et al. \(2018\)](#) that evaluated the same policy only using a geographically constrained RCT.

We continue by interacting our RD coefficients with measures of household head literacy, to show that large fractions of the total estimated effect of information on the receipt of *Raskin* rice is driven by household head literacy. Finally, leveraging administrative data detailing the precise locations and opening years of the universe of primary schools in Indonesia, we conduct a second policy evaluation that causally links school building to literacy gained during childhood. Specifically, we examine the role of graduating from primary school on literacy, while instrumenting having graduated with a variable that captures whether that individual lived in a village with an *SD INPRES* primary school during their primary school age.

Ultimately, we provide a chain of causal evaluations showing that the school building program dramatically affected household head literacy during childhood. Household head literacy in turn explains around half of the observed heterogeneity in the effect of the information campaign on household receipts of *Raskin* rice. Finally, we show that overall the information campaign resulted in a significant increase in poor household's receipt of *Raskin* rice. Taken collectively, our results highlight the prominent role of literacy in explaining the efficacy of information campaigns, an evaluation of which has been hampered due to the extraordinary data requirements.

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Tables

Table 1: The Effect of Information on Log (*Raskin* Bought) using RD Estimation

	Bandwidth: IK (2012)			Bandwidth: CCT (2014b)		
	All	Java	Non Java	All	Java	Non Java
	(1)	(2)	(3)	(4)	(5)	(6)
$E(R Information = 0)$ (Kg)	4.738	4.178	5.263	4.738	4.178	5.263
Panel A: Without Covariates-adjusted						
Linear	0.306 (0.146)	0.542 (0.445)	0.17 (0.156)	0.392 (0.131)	0.522 (0.386)	0.266 (0.144)
Quadratic	0.416 (0.13)	0.539 (0.385)	0.29 (0.142)	0.447 (0.126)	0.595 (0.341)	0.315 (0.134)
Cubic	0.457 (0.128)	0.611 (0.334)	0.328 (0.137)	0.402 (0.133)	0.759 (0.356)	0.243 (0.140)
Size of bandwidth [L: R]	[0.178: 0.198]	[0.168: 0.341]	[0.196: 0.198]	[0.115: 0.128]	[0.125: 0.229]	[0.129: 0.131]
Observations inside bandwidth	8,483	6,219	4,731	5,573	4,111	3,229
Observations	26,083	12,302	13,781	26,083	12,302	13,781
Panel B: With Covariates-adjusted						
Linear	0.259 (0.148)	0.568 (0.489)	0.14 (0.156)	0.35 (0.132)	0.503 (0.455)	0.225 (0.145)
Quadratic	0.381 (0.132)	0.513 (0.437)	0.255 (0.142)	0.41 (0.127)	0.495 (0.388)	0.270 (0.136)
Cubic	0.428 (0.129)	0.521 (0.404)	0.294 (0.139)	0.371 (0.135)	0.689 (0.394)	0.204 (0.143)
Size of bandwidth [L: R]	[0.180: 0.193]	[0.192: 0.419]	[0.202: 0.193]	[0.116: 0.124]	[0.129: 0.281]	[0.131: 0.127]
Observations inside bandwidth	8,322	7,435	4,676	5,496	4,971	3,174
Observations	26,083	12,302	13,781	26,083	12,302	13,781

Notes: This table displays nonparametric estimates of the effect of receiving information on the benefit received from the *Raskin* Program. The outcome variable is the log average *Raskin* rice bought in the last three months. All coefficients are estimated using a kernel local linear regression in an asymmetric bandwidth around the cutoff. $E(R|Z = 0)$ denotes the average monthly of *Raskin* Rice bought in the last three month by households who are not eligible for the KPS program ($Z = 0$). The table reports the bandwidth selection rule, IK2012 or CCT2014, the size of the bandwidth (distance from zero) and the number of observations included in the bandwidth. The standard errors (presented in parentheses) are clustered by the village.

Table 2: Kernel Local Linear Estimation at Selected Cut-Offs

Cutoffs	Bandwidth	Effect of Information	Robust Inference		Observation	
			<i>P-value</i>	<i>CI</i>	Left	Right
(1)	(2)	(3)	(4)	(5)	(6)	
-0.1	0.035	-0.01	0.947	[-0.200 : 0.187]	485	513
-0.05	0.026	-0.126	0.152	[-0.346 : 0.054]	444	460
0	0.177	0.071	0.013	[0.018 : 0.148]	2,670	5,106
0.05	0.037	0.024	0.47	[-0.095 : 0.260]	940	962
0.1	0.035	-0.05	0.335	[-0.218 : 0.074]	956	1,009

Notes: This table displays nonparametric estimates of the effect of receiving information on the benefit received from the *Raskin* Program at several different cut-offs. All coefficients are estimated using a kernel local linear regression in an asymmetric bandwidth around the cutoff. Optimal bandwidths are selected using IK2012. Robust *P – value* and *ConfidenceInterval* are reported in Column 4 and 5, respectively.

Table 3: The Effect of Information on Log (*Raskin* Bought) using LATE And LARF Estimations

	OLS	LATE			LARF		
	(1)	All Sample	Java	Non- Java	All Sample	Java	Non- Java
Reduced form		0.184 (0.003)	0.192 (0.006)	0.181 (0.004)			
Effect of Information	0.215 (0.012)	0.371 (0.043)	0.426 (0.066)	0.368 (-0.057)	0.485 (0.068)	0.465 (0.142)	0.426 (0.080)
First Stage Coef. of Z		0.226 (0.006)	0.217 (0.009)	0.238 (0.009)			
First Stage $F - Stat$ of Z		1239.46	598.1	687.95			
Control Village	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Control Vill. Head	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Observations	26,212	26,212	12,302	13,910	8,011	3,285	4,726

Notes: This table shows the estimates of the effect of receiving information on the benefit received from the *Raskin* Program. Dependent variables are the log average *Raskin* rice bought in the last three months. Column (1) is the estimation result using OLS estimation, ignoring the endogeneity on selection. The first stage instrument denotes a dummy $Z = 1$ if households are eligible, the first stage coefficient of Z and the F-statistic (for the excluded instrument which is adjusted for heteroskedastic and clustered standard errors) are also reported in Column (2) – (4). Column (2)-(4) is the LATE estimation result following Angrist, Imbens, and Rubin (1996). Column (5)-(7) is the LARF result following Abadie (2003). All standard errors are clustered at the village level and computed over the entire two-step using a block bootstrap with 500 repetitions following (Cameron, Gelbach and Miller, 2008).

Table 4: Information and Household Characteristics

	Dependent variable: Log (<i>Raskin</i> Bought)			
	(1)	(2)	(3)	(4)
Information	0.253 (0.022)	0.137 (0.064)	0.177 (0.035)	0.259 (0.024)
Information X <i>Head Can Read</i>		0.132 (0.067)		
Information X <i>Wife Can Read</i>			0.117 (0.043)	
Information X <i>Female Head of HHD</i>				-0.044 (0.050)
R^2	0.031	0.032	0.033	0.033
Clusters	3,958	3,958	3,958	3,958
Observations	12,854	12,854	12,854	12,854
Information effect (<i>Head Can Read</i>)		0.269 (0.023)		
Information effect (<i>Wife Can Read</i>)			0.293 (0.027)	
Information effect (<i>Female Head of HHD</i>)				0.217 (0.046)

Notes: Dependent variable is the log of *Raskin* bought. *Head Can Read* is indicator equal 1 if the head of the household can read, *Wife Can Read* is indicator equal 1 if the wife (for male-head household) can read, and *Female Head of HHD* is indicator equal 1 if head of the household is male. All standard errors are clustered at the village level.

Table 5: The Impact of Schooling on Head and Wife Literacy

	Head of Household			Wife		
	OLS		LATE	OLS		LATE
	(1)	(2)	(3)	(4)	(5)	(6)
Graduate from Primary School	0.353 (0.005)		0.494 (0.010)	0.708 (0.004)		0.733 (0.011)
First Stage Coef. of Z		0.185 (0.004)			0.209 (0.005)	
First Stage $F - Stat$ of Z		1764			1937	
Control Village	Yes	Yes	Yes	Yes	Yes	Yes
R^2	0.294		0.249	0.591		0.591
Observations	67,118	67,118	67,118	67,118	67,118	67,118

Notes: The dependent variable is literacy of head of household (male) and his wife. Column (1) & (4) are the estimation result using OLS estimation. Column (3) & (6) is the LATE estimation result following Angrist et al. (1996). Column (2) & (5) is the first stage estimation results including its F statistics of excluded instrument. All standard errors are cluster at village level.

Figures

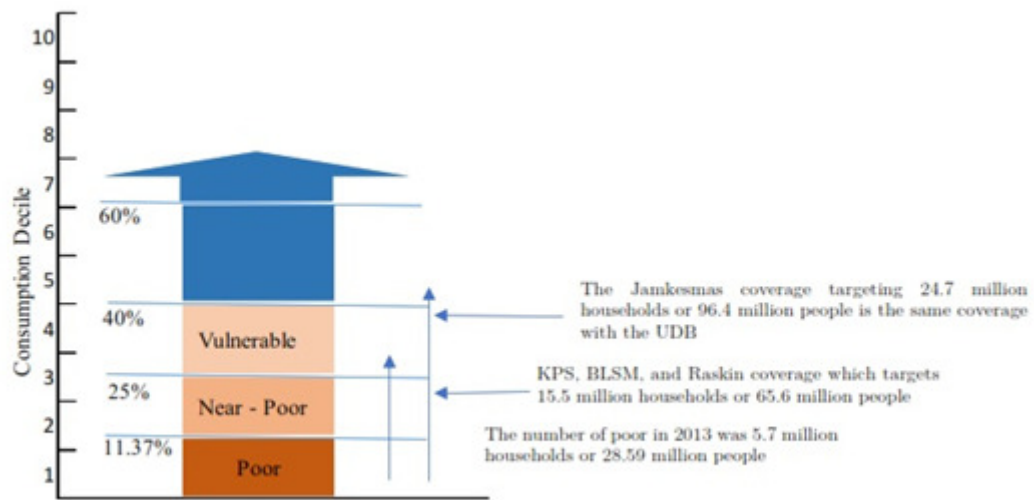


Figure 1: The UDB and the Coverage of Poverty Programs in Indonesia

Notes: This figure presents the coverage of the Unified Data and the biggest three poverty programs in Indonesia. For example in 2013, the percentage of the poor household is about 11.37, while the *Raskin* and BLSM covers approximately the bottom 25 percent of the population which the list of beneficiaries are extracted from the UDB.

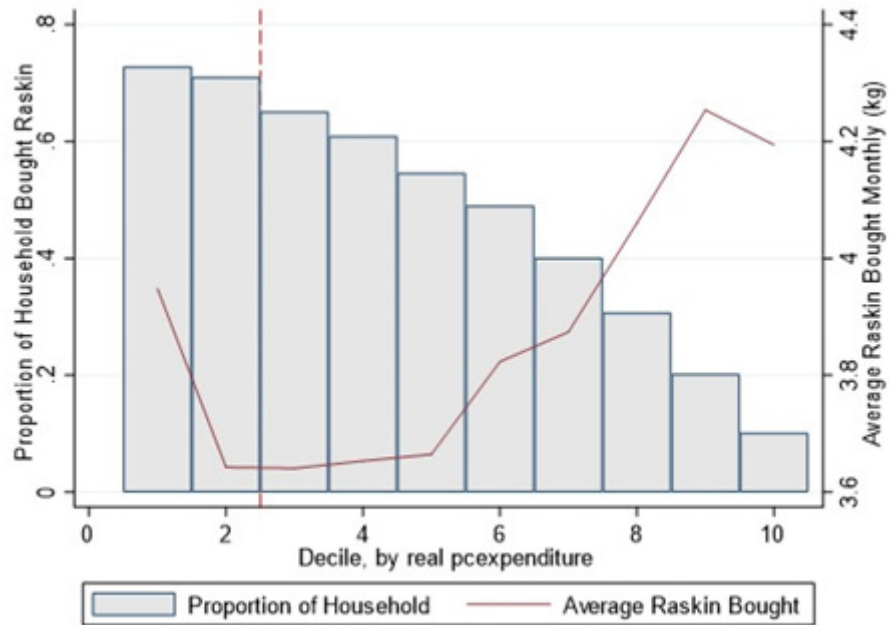


Figure 2: The Benefit Incidence of *Raskin* and the Average Rice Bought in each Decile, 2011

Notes: This figure shows the benefit incidence of *Raskin* program in 2011. The bar chart presents the percentage of the households in each decile who bought Rice. The line represents the average of rice (in kg) that was bought every month in each decile. The vertical dash line represents the threshold of the program. *Source:* Susenas 2011.

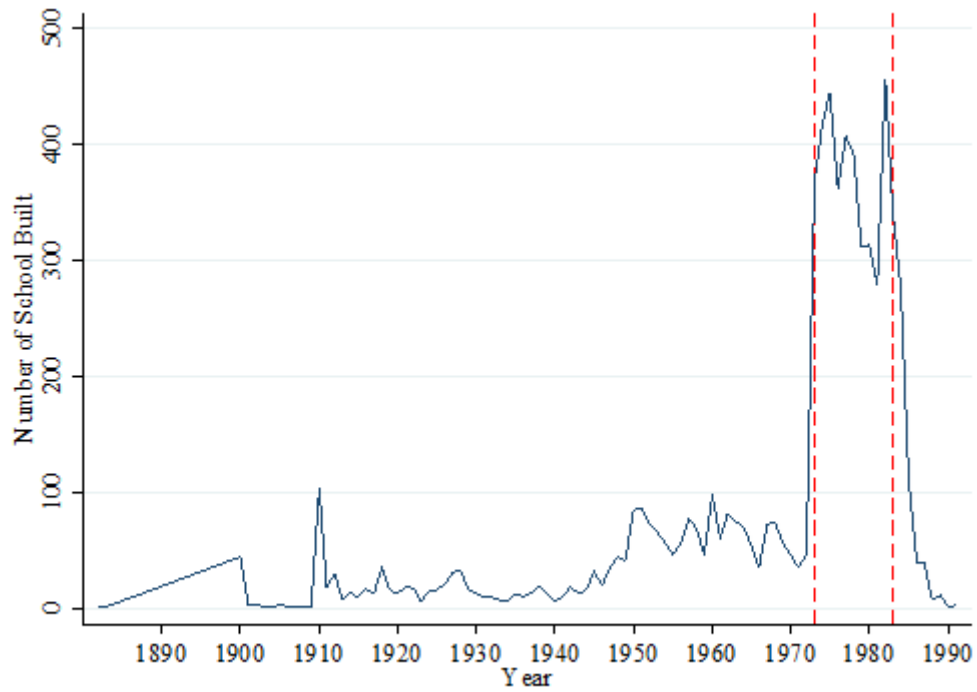


Figure 3: Number of Primary School Built in Indonesia From 1890 to 1990

Notes: This figure presents the number of primary school built for each year in Indonesia based on *DAPODIK* data. Within the vertical red dashed denotes the area in which the periods of *SD INPRES* program.

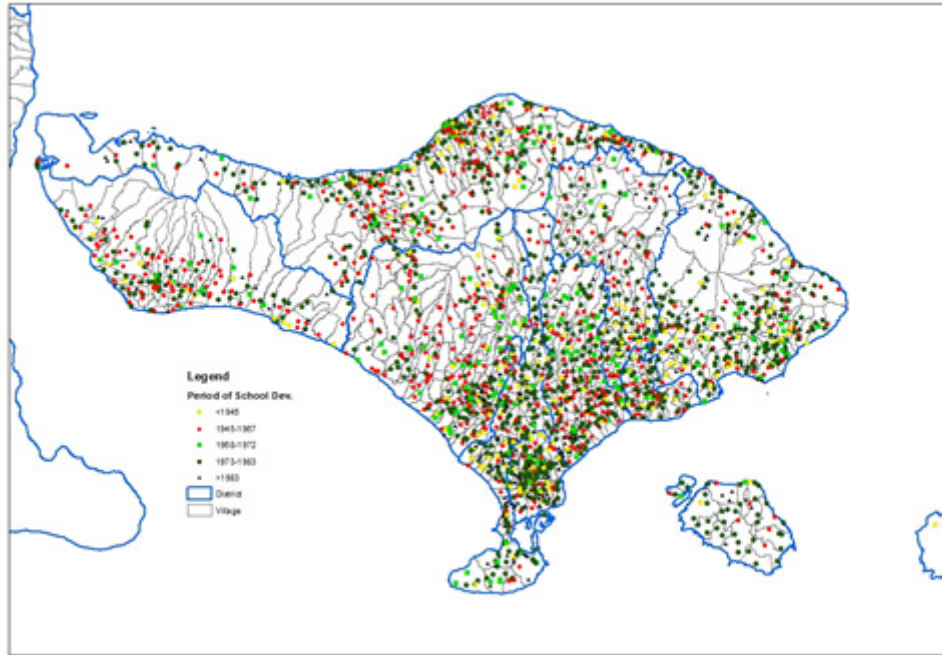


Figure 4: Illustration of School Location and Their Development Period in Bali Province

Notes: This figure shows the location of primary schools and their development periods in Bali, one of the Indonesian provinces. The blue line represents the district boundaries and the grey line is village or catchment area.

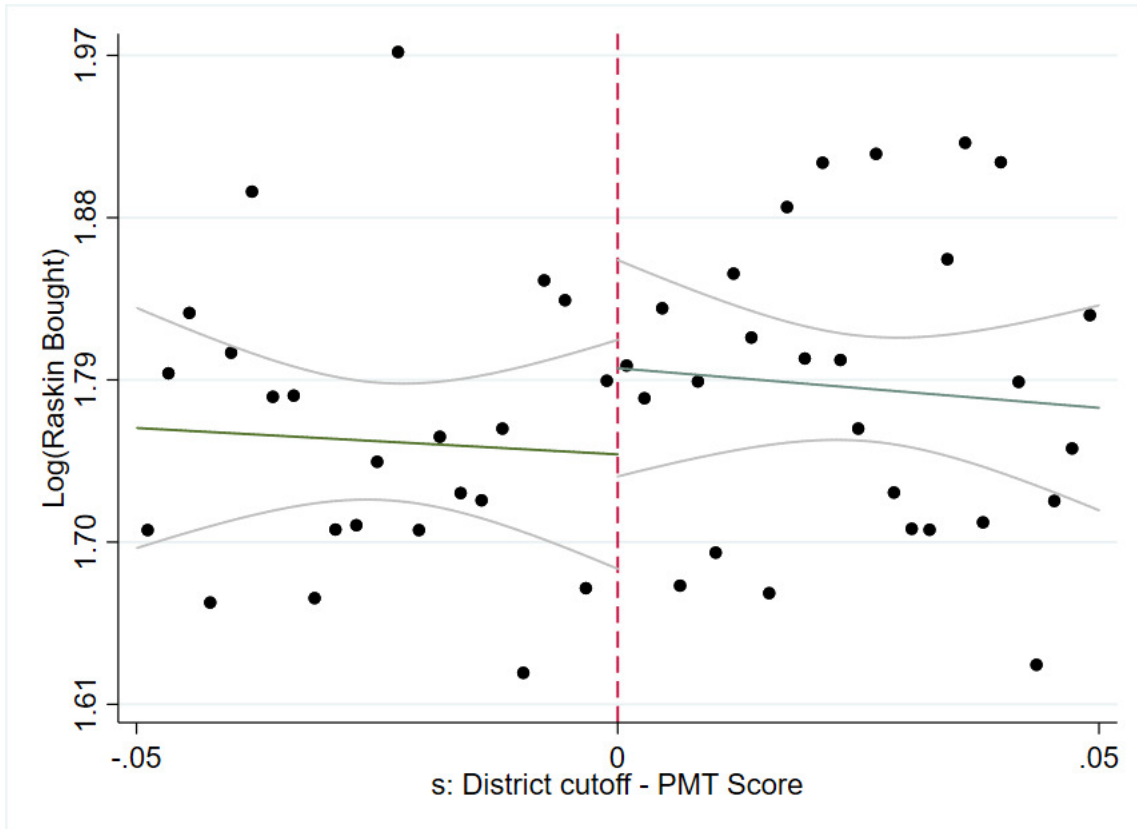


Figure 5: Discontinuity of Outcome variable at Cut-off ($s = 0$)

Notes: This figure represents graphical illustration of our RD design of $\text{Log}(\text{Raskin bought})$. The scatterplots are the average number within bins that are selected under IMSE-optimal quantile-spaced method using spacing estimators and the solid lines are the predicted outcomes, respectively. The bandwidth selection follows CCT2014.

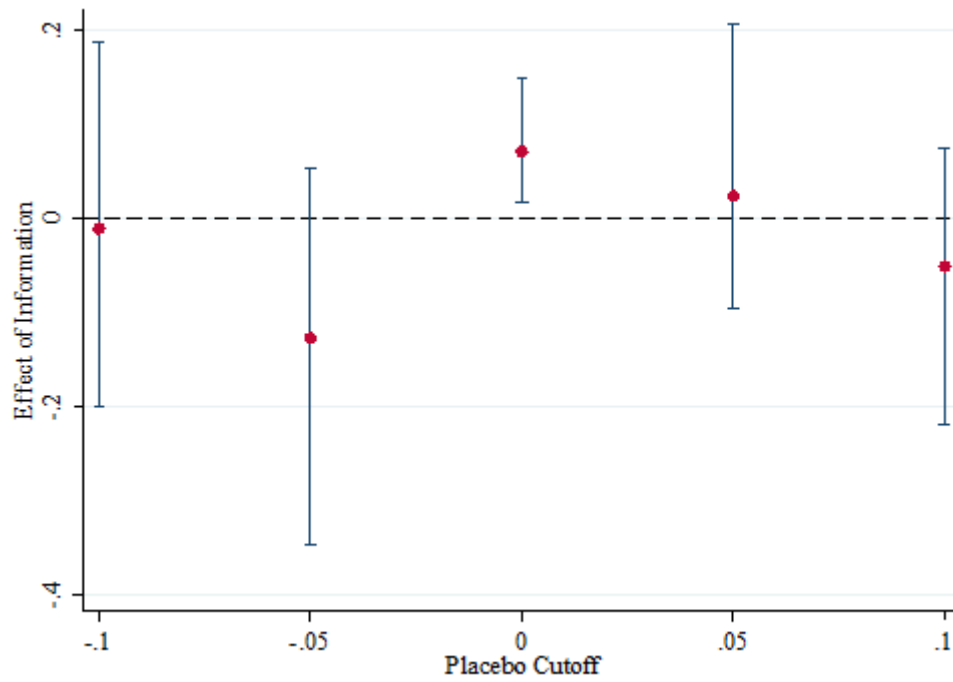


Figure 6: Sensitivity Analysis on Selected Cut-offs – All sample

Notes: This figure presents the sensitivity tests of the effect of information using different placebo cut-offs. The true cut-off, 0, is used as a benchmark for other artificial cut-offs. All coefficients are estimated using a kernel local linear regression in an asymmetric bandwidth around the cut-off. Optimal bandwidths are selected using IK2012.

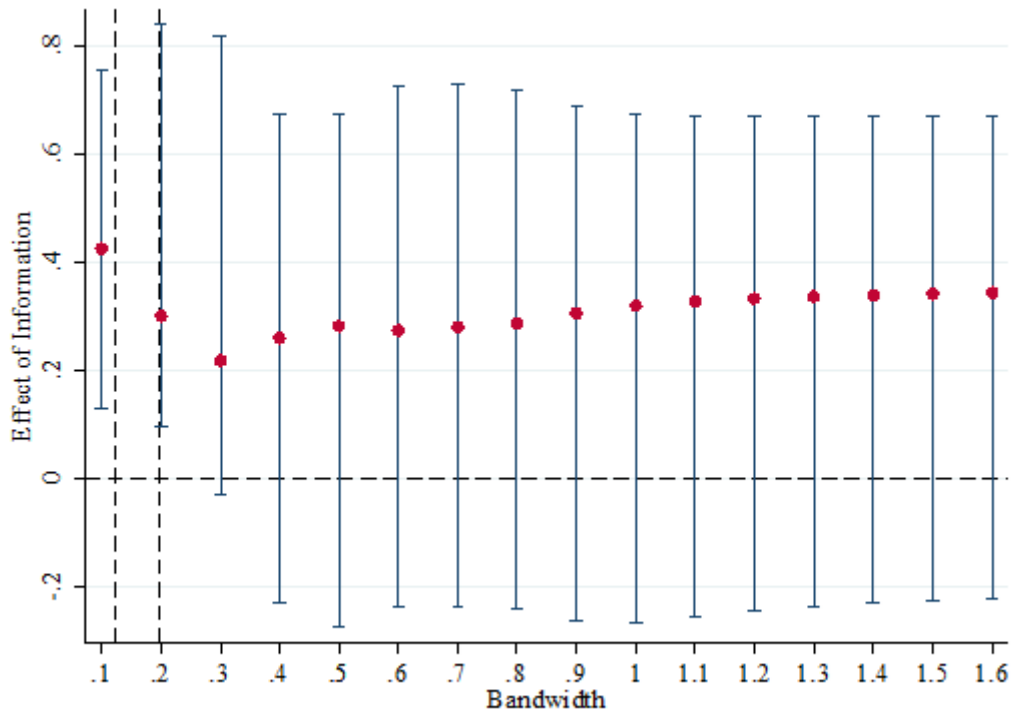


Figure 7: Sensitivity Analysis on Selected Bandwidths – All sample

Notes: This figure presents the sensitivity tests of the effect of information using different placebo bandwidth. Within the vertical dashed denotes the area in which optimal bandwidths are selected using IK2012 and CCT2014. All coefficients are estimated using a kernel local linear regression and blue lines represent the confidence intervals.

Appendix

A Tables

Table A.1: Outcome Variable and Household's Characteristics between Treatment and Control Groups of Raskin Beneficiaries

	Did Households receive informa- tion?				Difference	
	No		Yes			
	(1)	(2)	(3)	(4)	(5)	(6)
Monthly Raskin Bought (Kg)	4.738	(3.215)	6.012	(3.799)	1.274	[0.079]
Receive BLSM	0.145	(0.352)	0.969	(0.174)	0.824	[0.005]
PMT Score	13.462	(0.343)	13.298	(0.317)	-0.164	[0.006]
<i>Village Characteristics</i>						
Ln Distance to Nearest District office	2.914	(1.159)	2.818	(1.187)	-0.089	[0.026]
Ln Distance to Post office	1.651	(1.235)	1.642	(1.209)	-0.010	[0.029]
Availability of Asphalt Road in the village	0.752	(0.432)	0.76	(0.427)	0.008	[0.010]
Road can be accessed for a car	0.928	(0.258)	0.931	(0.254)	0.002	[0.007]
Cultural Mono	0.774	(0.418)	0.773	(0.419)	-0.001	[0.009]
Availability Access to the National TV Station	0.642	(0.479)	0.614	(0.487)	-0.028	[0.011]
Local Leader Directly Elected	0.84	(0.367)	0.81	(0.393)	-0.030	[0.008]
Sea Transport	0.037	(0.188)	0.034	(0.182)	-0.003	[0.004]
Padi as main Agriculture Product	0.49	(0.500)	0.5	(0.500)	0.009	[0.011]
Slum Area	0.094	(0.292)	0.093	(0.291)	-0.001	[0.006]
<i>Head of Village Characteristics</i>						
Male	0.933	(0.250)	0.922	(0.268)	-0.011	[0.006]
Age	44.437	(9.334)	44.173	(9.430)	-0.264	[0.204]
Education:						
No Education	0.013	(0.114)	0.01	(0.098)	-0.003	[0.003]
Primary	0.017	(0.131)	0.013	(0.111)	-0.005	[0.003]
Junior High	0.137	(0.344)	0.131	(0.338)	-0.006	[0.008]
Senior High	0.526	(0.499)	0.522	(0.500)	-0.004	[0.011]
University	0.045	(0.206)	0.048	(0.214)	0.004	[0.004]

Notes: Continue to the next page.....

	Did Households receive informa- tion?				Difference	
	No		Yes		(5)	(6)
	(1)	(2)	(3)	(4)		
<i>Head of Household Characteristics</i>						
Widow	0.151	(0.358)	0.151	(0.358)	0.000	[0.005]
Age	49.389	(13.892)	49.796	(13.547)	0.407	[0.209]
Years of schooling	6.319	(3.711)	5.519	(3.359)	-0.801	[0.055]
Position/Status of the main job:						
Self-Owned Business (SOB)	0.244	(0.430)	0.234	(0.423)	-0.010	[0.007]
SOB with non-permanent worker	0.262	(0.440)	0.259	(0.438)	-0.003	[0.008]
SOB with permanent worker	0.033	(0.179)	0.022	(0.148)	-0.011	[0.003]
Worker	0.347	(0.476)	0.373	(0.484)	0.026	[0.008]
Non Paid Worker	0.01	(0.099)	0.01	(0.101)	0.000	[0.001]
<i>Household Characteristics</i>						
Max years of schooling	8.974	(3.719)	8.381	(3.398)	-0.593	[0.056]
Dependency ratio	0.648	(0.643)	0.792	(0.692)	0.145	[0.010]
Urban area	0.338	(0.473)	0.334	(0.472)	-0.004	[0.010]
Receive the KPS from Postman	0.16	(0.367)	0.227	(0.419)	0.067	[0.021]
Number of households	19,032		7,180		26,212	

Notes: This table presents the averages of various outcome variables and household characteristics for treated and non-treated households. The numbers inside brackets represent standard deviations, while inside square brackets are standard errors.

Table A.2: The Impact of Schooling on Head and Wife Literacy - Migrant

			Head of Household			Wife		
			OLS		LATE	OLS		LATE
			(1)	(2)	(3)	(4)	(5)	(6)
Graduate from Primary School			0.265		0.318	0.902		0.757
			(0.009)		(0.016)	(0.003)		(0.017)
First Stage Coef. of Z				0.156			0.159	
				(0.006)			(0.007)	
First Stage F-Stat of Z				630.9			509.5	
Control Village			Yes	Yes	Yes	Yes	Yes	Yes
R2			0.235		0.226	0.805		0.785
Observations			18,997	18,997	18,997	28,475	28,475	28,475

Notes: The dependent variable is literacy of head of household (male) and his wife. Column (1) & (4) are the estimation result using OLS estimation. The estimations are conducted for only subgroup - Migrant. All specifications are the same as in [Table 5](#).

Table A.3: The Impact of Schooling on Head and Wife Literacy - Non Migrant

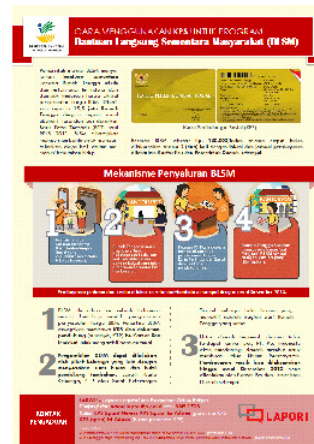
	Head of Household			Wife		
	OLS		LATE	OLS		LATE
	(1)	(2)	(3)	(4)	(5)	(6)
Graduate from Primary School	0.373 (0.006)		0.549 (0.012)	0.402 (0.007)		0.528 (0.014)
First Stage Coef. of Z		0.202 (0.005)			0.190 (0.006)	
First Stage F-Stat of Z		1438			999.9	
Control Village	Yes	Yes	Yes	Yes	Yes	Yes
R2	0.299		0.234	0.332		0.301
Observations	48,121	48,121	48,121	38,643	38,643	38,643

Notes: The dependent variable is literacy of head of household (male) and his wife. Column (1) & (4) are the estimation result using OLS estimation. The estimations are conducted for only subgroup - Non Migrant. All specifications are the same as in [Table 5](#).

B Figures



Panel A: Complaint mechanism of the KPS Card



Panel B: How to access BLSM program



Panel C: How to access Raskin program



Panel D: How to access Scholarship program

Figure B.1: Information included in the *KPS* package

Notes: The figures present the information included in the *KPS* package. Panel A is about complaint mechanism of the *KPS* in case the household has problem about their eligibility. Panels B, C, and D show the mechanism as to how *KPS* holders can access the benefit from BLSM, *Raskin*, and Scholarship programs respectively.



Front side of the KPS



Back side of the KPS

Figure B.2: The KPS Card

Notes: This figure shows the KPS card which includes information to protect the card from fraud.

Delivery Mechanism of *Raskin* Programs

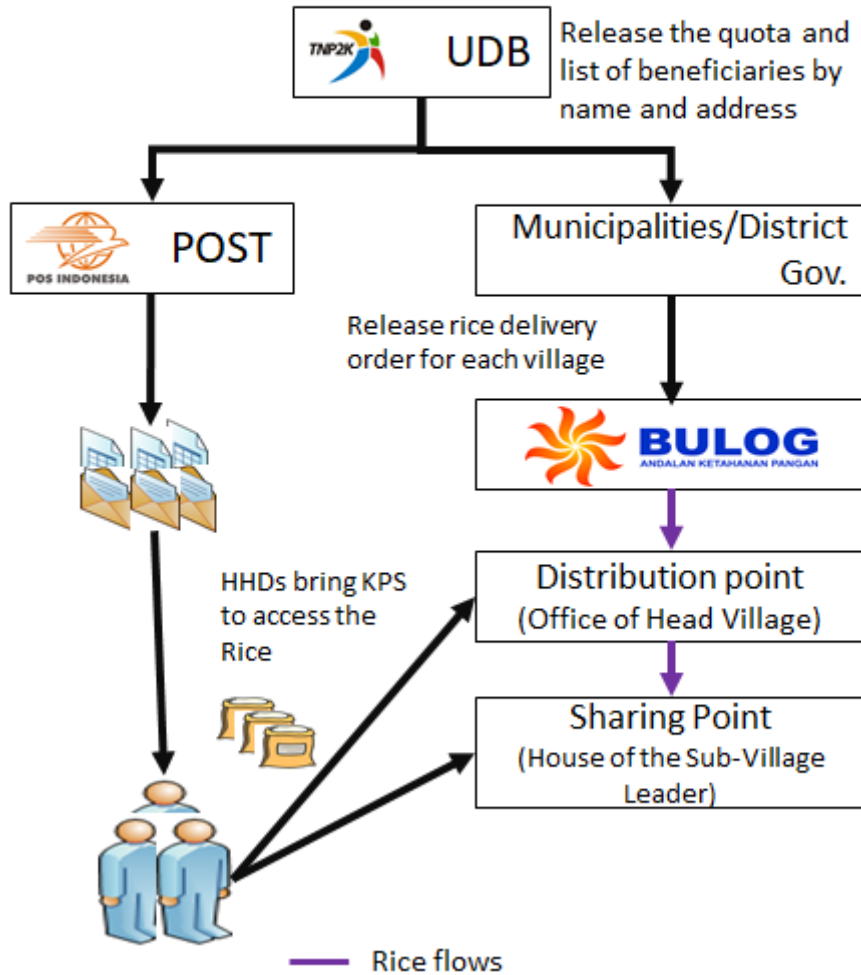


Figure B.3: The Delivery Mechanism of *Raskin* Programs

Notes: This figure shows the distribution of the *Raskin* rice relies on the authority of village leaders.

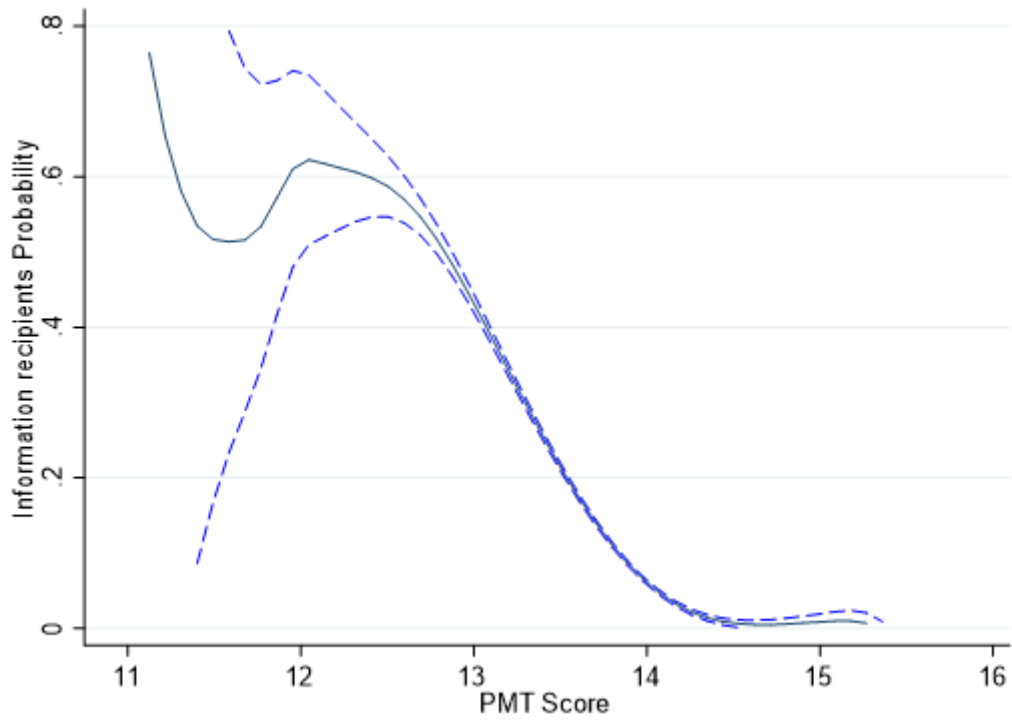


Figure B.4: KPS recipient versus PMT Score

Notes: This figure shows a nonparametric Fan regression of the estimated PMT Score against the probability of receiving information treatment. Bootstrapped pointwise 95 percent confidence intervals, clustered at the village level, are shown in dashes.