

Torres Gutierrez, Tania Paola; Ordóñez, Jessica Andrea

## Article

# Agglomeration economies and urban productivity

REGION

## Provided in Cooperation with:

European Regional Science Association (ERSA)

*Suggested Citation:* Torres Gutierrez, Tania Paola; Ordóñez, Jessica Andrea (2019) : Agglomeration economies and urban productivity, REGION, ISSN 2409-5370, European Regional Science Association (ERSA), Louvain-la-Neuve, Vol. 6, Iss. 1, pp. 17-24, <https://doi.org/10.18335/region.v6i1.242>

This Version is available at:

<https://hdl.handle.net/10419/235789>

### Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

### Terms of use:

*Documents in EconStor may be saved and copied for your personal and scholarly purposes.*

*You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.*

*If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.*



<https://creativecommons.org/licenses/by-nc/4.0>

## Agglomeration economies and urban productivity

Tania Torres-Gutiérrez<sup>1</sup>, Jessica Ordóñez<sup>1</sup>

<sup>1</sup> Universidad Técnica Particular de Loja, Loja, Ecuador

Received: 26 January 2018/Accepted: 21 November 2018

**Abstract.** This study explores the relationship between agglomeration economies and industrial productivity between 1980 and 2010 in Ecuador. The measure of productivity used is labor productivity. We conclude that urbanization economies have a positive impact on productivity in the period analyzed. These results are consistent with other works for developed and developing countries.

**JEL classification:** R12

**Key words:** agglomeration, agglomeration economies, urban productivity, Ecuador

### 1 Introduction

The increasing concentration of people and production produces benefits known as economies of agglomeration in the economic literature. Traditionally, agglomeration economies are classified in location economies and urbanization economies.

From the seminal works of Glaeser et al. (1992) and Henderson et al. (1995), the ongoing debate is not only about the dichotomy between specialized and diversified environments within the same urban system, but also about the coexistence of specialization and diversity.

Following Glaeser et al. (1992), location economies or MAR<sup>1</sup> externalities that operate within a specific industry restrict the flow of ideas to others, allowing the innovator to internalize externalities. Such interactions can positively influence the productivity of companies and the growth of cities. On the other hand, the urbanization of economies occurs through industries, which motivates the argument of Jacobs (1969) that the variety of industries within a geographical region promotes knowledge spillovers and results in innovative activities and economic growth. In this framework, the concepts of specialization and diversification are inherent to the economies of location and urbanization, respectively.

The empirical literature establishes that spatial concentration of industrial activity improves economic growth, productivity, and innovation through different approaches, among which the common denominator is the analysis of the location-urbanization dichotomy. In line with this literature, this study explores the relationship between the economies of agglomeration and industrial productivity between 1980 and 2010, years for which census data exist for the economic activity of the country. There are two motivations for this work. First, to contrast the economic literature and the empirical results broadly focused in developed countries with those of a developing country like

<sup>1</sup>Refers to the model presented by Marshall (1890), Arrow (1962), and Romer (1986).



Source: Own elaboration based on INEC data

Figure 1: Average annual growth of productivity in the cantons of Ecuador: 1980 – 2010

Ecuador. Second, to contribute to the orientation and reform of economic policies related to the productivity of the country, which seeks to reorient its strong agro-export sector.

## 2 Industrial dynamism: Diversity of the cantons of Ecuador

The industrial sector is the second most important in terms of GDP in the Ecuadorian economy, but it is the most dynamic given that, unlike the other sectors, it has experienced 9% growth between 1980 and 2010 according to World Bank data. The services sector is the most relevant due to, among other things, the momentum generated in the eighties by exports and the oil boom that stimulated this sector, as well as that of the public administration. This is compounded by the significant growth in self-employed activities in the tertiary sector, whose participation in the national economically active population (EAP) in 1974 was 8.4%, 11.1% in 1982, and 28.5% in 2010. Finally, agriculture has fallen in its share of the national GDP by 6% in this period.

As in other countries, economic activity tends to agglomerate in relatively few cities: Guayaquil and Quito mainly, those that from colonization maintain their supremacy over the others, and therefore perform important economic, regional, and international functions. Although these cities have altogether only 3.29% of the total surface area, they contain 16.25% and 15.48% of the population, they generate 23.61% and 25.19% of the gross added value, and represent 21.35% and 28.91% of manufacturing employment, respectively. Based on the information available for Ecuador at the industry and canton level, Figure 1 shows the productivity growth in the analysis period for cantons whose increase is above the average annual growth rate of 4.2%.

## 3 Data and variables

To determine if agglomeration economies affect productivity, we used data from the 1980 and 2010 Economic Census of the National Institute of Statistics and Censuses (INEC), at the level of each sector and canton or municipality, except for those corresponding to the Galapagos Islands. The empirical work included the homogenization of the databases prepared from the referred censuses because these were not directly comparable. In total,

Table 1: Summary statistics of the main variables

Variable	Mean	Standard Deviation
productivity growth	.127	1.391
specialization	-.097	.815
diversity	-.983	.470
density	-1.711	1.958
average size firms	-.2509	.609

Notes: The productivity growth is between 1980 and 2010. All variables refer to logarithms

26 sectors and 114 cantons are integrated. The sectoral breakdown corresponds to two digits and three digits - ISIC for 1980 and 2010, respectively.

The dependent variable is measured as follows:

$$\Delta prod_{ic} = \frac{\log(Y_{ic-2010}/emp_{ic-2010})/(Y_{i-2010}/emp_{i-2010})}{\log(Y_{ic-1980}/emp_{ic-1980})/(Y_{i-1980}/emp_{i-1980})} \quad (1)$$

where,  $Y_{ic}$  and  $emp_{ic}$  are the levels of production and employment by industry and canton, respectively, between 1980 and 2010.

### 3.1 Measuring agglomeration economies

The measure of agglomeration economies is the index of specialization related to location economies:

$$esp_{ic} = \frac{emp_{ic}/emp_c}{emp_i/emp} \quad (2)$$

where,  $emp_{ic}$  is the employment of industry  $i$  in canton  $c$ , and  $emp$  is total employment.

While the economies of urbanization are commonly measured through the inverse of the Herfindahl index, constructed from the participation of industries in local employment, with the exception of the industry that is considered, this variable is normalized by the same variable at the country level:

$$div_{ic} = \frac{1/\sum_{i*=1, i* \neq i}^i [emp_{i*c}/(emp_c - emp_{ic})]^2}{1/\sum_{i*=1, i* \neq i}^i [emp_{i*}/(emp - emp_i)]^2} \quad (3)$$

where  $i$  is the number of industries. The numerator is maximum when all sectors, except the subject of the analysis,  $i^*$ , are the same size in the cities. This indicator reflects the sectorial diversity of the industry and the city. Therefore, it is not necessarily related to the level of specialization of the industry being analyzed.

With the intuition that large companies are usually better able than small companies to internalize some of the local effects, Glaeser et al. (1992) suggest incorporating the average size of firms within the local industry as an additional determinant of location economies. When normalized by the average of the companies in the industry at the level of the whole country, we obtain:

$$size_{ic} = \frac{emp_{ic}/n_{ic}}{emp_i/n_i} \quad (4)$$

where  $n_{ic}$  is the number of companies in the industry and in city  $c$ . However, according to Combes, Gobillon (2015), its use leads to serious problems of endogeneity, since it depends on the location options of the companies and their scale of production, which directly influence local productivity. Thus, one should avoid introducing it into the specification unless you have a strong instrumentation strategy.

Finally, as in Combes (2000), to simultaneously control for differences between cities, it is relevant to consider the density of total employment by means of the following indicator:

$$den_c = \frac{emp_c}{area_c} \quad (5)$$

where  $area_c$  is the area of the city measured in  $\text{km}^2$ .

### 3.2 Selection bias

A particularity of the data used in this study is that they are not fully observable, since some industrial sectors are present only in certain cities. This is a typical problem in research that uses data on a local scale. We isolate the selection bias through a model Heckman (1979) proposes, applied in two stages as in Viladecans-Marsal (2004), Combes (2000), Henderson et al. (1995), and others. In the first stage, a model is formulated to estimate the probability that a city contains an industrial sector:

$$Prob(S = 1|Z) = \Phi(Z\gamma) \quad (6)$$

where  $S$  indicates the sector ( $S = 1$  if the sector is in the city and  $S = 0$  otherwise),  $Z$  is a vector of explanatory variables,  $\gamma$  is a vector of unknown parameters, and  $\Phi$  is the cumulative distribution function of the normal distribution. The estimation of the model yields results that can be used to predict the probability that an industry is contained in a specific city. In the second stage, the initial model is estimated by OLS with the dependent variable of continuous productivity growth, which corrects the selection bias by incorporating the variable called the Mills inverse ratio ( $\lambda$ ), which is derived from the previous stage.

### 3.3 Endogeneity and instrumental variables

When estimating the impact of agglomeration economies on local results such as productivity growth, the literature recognizes two potential sources of endogeneity: omitted variables and inverse causality. Either may arise at the local and individual level (Combes, Gobillon 2015), and their treatment focuses on instrumental variables, including historical and geographical variables, for each endogenous regressor, specialization, diversity, and density.

According to Combes, Gobillon (2015), historical values of population or density are relevant, because by remaining in time, they create inertia in the population and in local economic activity. This idea is imputed to the construction of the instruments for the variables of specialization and diversity. For that reason, the instruments are generated from the data of the birth of firms. The birth of firms is considered correlated with the level of specialization and diversity of the industries in the cities, but not directly with the growth of employment and productivity. The dummy of geological character is related to all the endogenous variables and does not represent a direct effect on the variables of interest or the geographical variables related to the availability of roads in 1980, also generated as dummies.

Estimating the effect of location, urbanization economies, and density on productivity using instrumental variables can lead to unbiased estimates, provided that the instruments meet the conditions of relevance (7) and exogeneity (8). Formally, these conditions are:

$$\begin{aligned} Cov(Specialization_a, Z_a|.) &\neq 0 \\ Cov(Diversity_a, Z_a|.) &\neq 0 \\ Cov(Density_a, Z_a|.) &\neq 0 \end{aligned} \quad (7)$$

$$\begin{aligned} Cov(\mu_a^x, Z_a) &= 0 \\ \text{for } x &= \text{productivity} \end{aligned} \quad (8)$$

$Z$  denotes the set of instruments.

Equation (7) denotes that the relevance of an instrument depends on the partial correlation of the instrumental variables and the endogenous regressors. These are obtained by Ordinary Least Squares (OLS) estimates for each endogenous regressor of the growth of productivity in instrumental variables, including regional, provincial, and industrial fixed effects. The results show that the relevant instruments for specialization are the specialization index, the population density of 1950, the urbanization index of 1980, the availability of roads in 1980, and a geological dummy. For the diversity and density of employment these same instruments are relevant, except the specialization index.

The analysis of the relevance of the defined instruments is validated by the test developed by Stock, Yogo (2005)<sup>2</sup>, who define two tests for weak instruments based on a single  $F$  statistic. The values in all cases are greater than 10<sup>3</sup>, suggesting that the instruments are strong; their strength is confirmed when they contradict the critical values reported by Stock, Yogo (2005).

The condition of exogeneity suggested in equation (8), that is, the orthogonality with respect to the error term, is evaluated with the Sargan over-identification test, which allows us to reject the hypothesis of restriction of over-identification, suggesting the joint exogeneity of the instruments.

#### 4 Estimation and analysis

The model specified to estimate the effect of agglomeration economies on the productivity growth of a particular industry in a certain canton between 1980 and 2010 appears below. In particular, the Two-Stage Least Squares (2SLS) estimator is used due to the aforementioned aspects of endogeneity, focused on regressions with instrumental variables.

$$\Delta prod_{ic} = \beta_0 + \beta_1 \log spe_{ic} + \beta_2 \log div_{ic} + \beta_3 size_{ic} + \beta_4 den_c + imr_{-e} + imr_{-p} + \varepsilon_{ic} \quad (9)$$

where  $\Delta prod_{ic}$  represents the growth of productivity in industry  $i$  and city  $c$  between both years, respectively;  $spe_{ic}$ ,  $div_{ic}$ ,  $size_{ic}$  are the indices of specialization, diversity, and average industry size  $i$  in city  $c$ ; and  $den_c$  is the density of total employment. The variables  $imr_{-e}$  and  $imr_{-p}$  are the inverse ratio of Mills for employment and productivity in each case, introduced to control the selection bias, and  $\varepsilon_{ic}$  is assumed as the error term. To control for unobservable heterogeneity, we introduce fixed effects at the province, industry, and regional<sup>4</sup> levels. The literacy rate of each city aims to capture the qualifications of the population in each case<sup>5</sup>.

The explanatory variables correspond to the initial year, 1980, and have been normalized by the corresponding values at the national level. All the variables are expressed in logarithms, which is why the estimated parameters are their elasticities with respect to each variable. This makes them easily comparable and interpretable.

The first estimates of equation (9) are made by OLS. However, given the presence of selection bias and the endogeneity of the model, such results are not entirely correct, as it is pertinent to apply two additional estimation strategies. To correct the selection bias, we proceeded with maximum likelihood estimations through a Tobit Type II model, while the endogeneity of the model implies estimations with instrumental variables (2SLS) with results accepted as definitive (Table 2).

As a robustness test, a strategy for estimating productivity growth is applied that is less sensitive to weak instruments: the limited information maximum likelihood (LIML)

<sup>2</sup>Stock, Yogo (2005) provide two tests that, based on the  $F$  statistic, have two purposes. The first is to test the hypothesis that in small samples the bias in the 2SLS regressions is small with respect to the endogeneity bias reported by MCO ("bias test"). The second is to use the Wald test to determine whether an instrument is considered strong, that is, that its size is close to its level for all possible configurations of the regression by instrumental variables ("size test"). Therefore, the instruments may be weak in one sense but not in another.

<sup>3</sup>Cameron, Trivedi (2010) indicate that a measure widely used by Staiger, Stock (1997), that is,  $F < 10$ , suggests weak instruments.

<sup>4</sup>In Ecuador, there are three natural regions: Coast, Highland, and Amazonian regions.

<sup>5</sup>No data is available on the qualification of employees by industry and city.

Table 2: Productivity growth: Estimates by OLS, Tobit and IV

Productivity	OLS	TOBIT	2SLS
specialization	-0.075* (0.026)	0.091 (0.073)	-0.285 (0.166)
diversity	0.064 (0.046)	-0.189 (0.178)	0.397* (0.183)
density	0.031 (0.019)	-0.019 (0.043)	-0.161* (0.081)
size firms	-0.021 (0.036)	0.229* (0.101)	0.049 (0.077)
inverse Mills ratio	-0.032 (0.119)		-0.767 (0.341)
N	2963	2963	2963
F	—	—	14.75
CONTROL			
Literacy index 1980	Yes	—	Yes
FIXED EFFECTS			
Region	Yes	No	Yes
Province	Yes	No	Yes
Industry	Yes	Yes	Yes
Over identification (Sargan Test)	—	—	1.145
P value - SarganTest	—	—	(0.5640)

Notes: Standard errors in parentheses. P-values: \*  $p < 0.1$ , \*\*  $p < 0.05$ , \*\*\*  $p < 0.01$ . All variables are expressed in logarithms.

estimator. This strategy takes into account only the likelihood function of the endogenous variables of equation (9) and the identification of restrictions corresponding to the equation to be estimated.

The results obtained are supported with those obtained by 2SLS and are consistent with previous empirical findings given that the productivity gains of urban agglomeration economies are generally found to be positive (Melo et al. 2009). The results obtained differ in that the location economies are not significant. At this point, it should be noted that 94% of the firms of the two years analyzed correspond to the category of microenterprise, 4% correspond to small firms, and the 2% remaining percentage are medium and large firms<sup>6</sup>. This corresponds to the finding of Jacobs (1969) that small businesses benefit more from urban diversity in large cities due to their greater dependence on external industrial environments for multiple intermediate inputs, while large companies are self-sufficient.

## 5 Conclusion

In particular, two different contributions to the literature are presented. The first relates to the agglomeration literature about Latin American countries like Ecuador that have received little attention from this approach.

The growth of productivity is determined significantly and positively by urbanization economies, while the density elasticity is negative. This is interpreted as the result of the effects of congestion. These results are consistent with other works for developed countries Cingano, Schivardi (2004) and Guevara et al. (2015) for Ecuador. They approximate labor productivity in 2010, both in industry and in services, as a function of specialization, diversity, competence, and density (of firms or employment) in the cities of Ecuador. Using as main instruments the spatial delays of each of the endogenous variables, their results suggest a strong positive externality of the diversity in the productivity of the

<sup>6</sup>In Ecuador, companies are classified as micro, small, medium, and large depending on whether they have between 1 and 9, 10-49, 50-199, or more than 200 employees, respectively.

manufacturing industries (1,651) and of the services (2,081). In manufacturing, the density of employment is also statistically significant.

Productivity takes place in provincial capitals, characterized by the concentration of public sector intervention, both in terms of investments (public goods) and public consumption (services); ease of access to large markets; and the possibility of finding large niches of specialization, and access to a broad labor market and specialized urban functions.

Second, in terms of Ecuador's public policy, a boost to industry and services is expected within the framework of the country's industrial policy. For the period 2016-2025, this policy aims to generate 251,000 new jobs; to invest 13,600 million dollars; contribute positively to the trade balance of 1,200 million dollars; and increase GDP by 10 percentage points. Ecuador's public policy recognizes the imminent change in the spatial distribution of economic activities and that it is necessary to focus on land use and labor mobility relating to trade in intermediate goods. Consequently, the industrial policy must mesh with others that consider the spatial dimension. In the area of externalities in particular, the challenge is to balance negative externalities and exploit the positive externalities of agglomeration based on greater knowledge, an important mechanism through which the agglomeration economies act.



## References

- Arrow K (1962) The economic implications of learning by doing. *Review of Economics and Statistics* 29: 155–173
- Cameron AC, Trivedi PK (2010) *Microeconometrics Using Stata* (Revised ed.). Stata Press, College Station, TX
- Cingano F, Schivardi F (2004) Identifying the sources of local productivity growth. *Journal of the European Economic Association* 2[4]: 720–744. [CrossRef](#).
- Combes P, Gobillon L (2015) The empirics of agglomeration economies. In: Duranton G, Henderson V, Strange W (eds), *Handbook of Regional and Urban Economics*, Volume 5. North Holland, Amsterdam, 247–348
- Combes PP (2000) Economic structure and local growth: France, 1984–1993. *Journal of urban economics* 47[3]: 329–355
- Glaeser E, Kallal H, Scheinkman J, Shleifer A (1992) Growth in cities. *Journal of Political Economy* 100[6]: 1126–1152. [CrossRef](#).
- Guevara C, Riou S, Autant-Bernard C (2015) Agglomeration externalities and urbanization in Ecuador. 55th congress of the European Regional Science Association, Lisbon, Portugal, <https://halshs.archives-ouvertes.fr/halshs-01212174>
- Heckman J (1979) Sample selection bias as a specification error. *Econometrica* 47[1]: 153–161
- Henderson J, Kuncoro A, Turner M (1995) Industrial development in cities. *Journal of Political Economy* 103[5]: 1067–1090. [CrossRef](#).
- Jacobs J (1969) *The Economy of Cities*. Random House, New York
- Marshall A (1890) *Principles of Economic*. MacMillan, London
- Melo PC, Graham DJ, Noland RB (2009) A meta-analysis of estimates of urban agglomeration economies. *Regional Science and Urban Economics* 39[3]: 332–342. [CrossRef](#).
- Romer P (1986) Increasing returns and long-run growth. *Journal of Political Economy* 94: 1002–1037
- Staiger D, Stock J (1997) Instrumental variables regression with weak instruments. *Econometrica* 65: 557–586
- Stock J, Yogo M (2005) Testing for weak instruments in linear iv regression. In: Andrews D, Stock J (eds), *Identification and Inference in Econometric Models: Essays in Honor of Thomas J. Rothenberg*. Cambridge University Press, Cambridge, UK, Chapter 5
- Viladecans-Marsal E (2004) Agglomeration economies and industrial location: City-level evidence. *Journal of Economic Geography* 4[5]: 565–582. [CrossRef](#).

