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Article

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Review of Economic Perspectives

Provided in Cooperation with:

Masaryk University, Faculty of Economics and Administration

Suggested Citation: Vasilev, Aleksandar (2020) : Are "fair" wages quantitatively important for business cycle fluctuations in Bulgaria?, Review of Economic Perspectives, ISSN 1804-1663, De Gruyter, Warsaw, Vol. 20, Iss. 1, pp. 91-105, <https://doi.org/10.2478/revecp-2020-0005>

This Version is available at:

<https://hdl.handle.net/10419/227542>

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Are “fair” wages quantitatively important for business cycle fluctuations in Bulgaria?

Aleksandar Vasilev¹

Abstract: We introduce “fair” wages in a general-equilibrium model where worker’s effort is unobservable and investigate whether such a mechanism can quantitatively account for the degree of real wage rigidity in the Bulgarian labor markets, as documented in Lozev, Vladova, and Paskaleva (2011) and Paskaleva (2016). In contrast to Danthine and Kurmann (2004), we internalize the effect that past wages have on the current effort level. We calibrate the model to Bulgarian data (1999–2016), and quantify the effect of technological shocks on hours and wages in the theoretical setup. Overall, the calibrated model with “fair” wages performs poorly when it comes to the relative volatilities of labor market variables. This is because aggregate labor market conditions, as proxied by the employment rate and past aggregate wages, turn out not to be quantitatively important for business cycles in Bulgaria.

Keywords: Business cycles, unobservable effort, fair wages, unemployment, Bulgaria

JEL Classification: E24, E32, J41

Received: 3 April 2019 / Accepted: 10 December 2019 / Sent for Publication: 12 March 2020

Introduction and Motivation

In this paper, we investigate the quantitative importance of unobservable worker’s effort, and the downward-rigid efficiency (“fair”) wages that are introduced as a result of that informational imperfection, in explaining business cycle fluctuations in Bulgaria after the introduction of the currency board arrangement in 1997. Earlier macroeconomic literature, using Dynamic Stochastic General Equilibrium (DSGE) models with perfectly-competitive labor markets, e.g., Vasilev (2009), was not able to capture well the dynamics exhibited by labor market variables (wages, employment, and unemployment) in Bulgaria. That is why we adapt the standard model, and augment it with a plausible mechanism in labor markets that deviates from spot wage contracting, and instead move to setups that emphasize the long-term aspects of the labor arrangement. After all, the employer-employee relationship is a multi-period contract problem. Therefore, an alternative mechanism of wage contracting is considered here, as those mechanisms, mostly

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based on non-Walrasian settings, are promising areas of research, as pointed in Vasilev (2017c). Similarly, including a government sector, in addition to making the model more realistic, helps the standard Real-Business-Cycle model match data better.

As shown in Paskaleva (2016), real wages in Bulgaria are indeed downward rigid. That is mostly due to collective agreements in place, which prohibit cuts in base wages. Such restrictions mean that adjusting labor costs needs to happen mostly through employment reductions. Lozev, Vladova, and Paskaleva (2011) also document downward real wage rigidity in Bulgaria, even though it is lower than in the other EU member states. We use these empirical findings to motivate our modelling approach here. In contrast to Vasilev (2017b) who introduces wages of no-shirking type a la Shapiro and Stiglitz (1984), and Vasilev (2017c), who incorporated reciprocity arrangement in labor relations and wage determination a la "gift exchange" as in Akerloff (1982), into a relatively standard RBC model with government sector, here we follow Danthine and Kurmann (2004). As in Vasilev (2017b,c), effort will be modelled as a productive input in the firm's production function, but it will be unobservable from the employer's perspective, and thus a contract fully specifying the required level of effort can be neither specified nor enforced. The novelty in this paper is that conditional on working, each household suffers additional dis-utility from exerting effort on the job. However, they will be willing to supply an effort level beyond and above some norm if they feel they have been treated in a fair manner by their employer. The extra effort would come as reciprocation for a wage rate w_t above some reference level. In addition, the effort function depends on the overall labor market conditions, the going wage, and past compensation. Allowing effort to depend on past wages helps to introduce sluggishness in wage adjustments in the theoretical framework, and that is what the standard gift exchange model, e.g., Akerlof (1982) lacks.

In addition, Collard and de la Croix (2000) and Danthine and Kurmann (2004) show that including past wage considerations in the reference wage function helps a general-equilibrium macroeconomic model generate substantial downward wage rigidity and improves the overall statistical fit. In contrast to Danthine and Kurmann (2010), here we stay within the RBC paradigm, and instead of estimating the model, we calibrate it to Bulgarian data for the period 1999-2016, which corresponds to the period of stability after the introduction of the currency board arrangement. In addition, for better realism, we introduce a detailed government sector and analyze the business cycle properties of the model relative to the data in much more detail as compared to earlier studies.

The rest of the paper is organized as follows. Section 2 describes the model. Section 3 described the data used and the calibration procedure. Section 4 presents the long-run theoretical properties of the model. Section 5 simulates the model and evaluates its business cycle properties vis-a-vis data in the spirit of Canova (2007), especially the response of main variables to an unanticipated technology shock. Section 6 concludes.

Model Setup

Description of the model

There is a continuum of identical one-member households distributed on the [0, 1] interval and indexed by i . Each household i derives utility out of consumption and leisure. As in Vasilev (2018), each household i maximizes the following utility function:

$$E_0 \sum_{t=0}^{\infty} \beta^t \{ \ln c_t^i - h_t^i G(e_t^i) \} \tag{2.1}$$

Where E_0 is the expectation operator conditional on the information available to household i at time 0, $0 < \beta < 1$ denotes the common to all households discount factor, c_t^i is the consumption of household i in period t , h_t^i is the fraction of time available to household i that is spent working, and e_t^i is the level of effort exerted. The total time endowment available to each household is normalized to unity, and $l_t = 1 - h_t$ is implicitly expressed as time off work.

The novelty in this relatively standard utility function is the last term. In particular, as in Vasilev (2018), conditional on working, each household suffers additional dis-utility from exerting effort on the job. However, they will be willing to supply an effort level beyond and above some norm if they feel they have been treated in a fair manner by their employer. The extra effort would come as reciprocation for a wage w_t above some reference level. In addition, the effort function depends on the overall labor market conditions, the going wage, and past compensation – all those factors are empirically supported in surveys, e.g., Bewley (1998).

Following Collard and de la Croix (2000) and Danthine and Kurmann (2007), we express the effort function as

$$G(e_t^i) = [e_t^i - (\phi_0 + \phi_1 \ln w_t^i + \phi_2 \ln n_t + \phi_3 \ln w_t + \phi_4 \ln w_{t-1}^i)]^2 \tag{2.2}$$

Where $\phi_0, \phi_1 > 0, \phi_2, \phi_3, \phi_4 < 0, \phi_1 + \phi_3 > 0, w_t^i$ is the household’s wage in period t , and w_t denotes the average wage in the economy.

As in Hansen (1985) and Rogerson (1988) household i ’s labor supply is assumed to be indivisible. In equilibrium, as demonstrated in Vasilev (2018), only a fraction n_t would be selected to work a full shift in each period t . In order to Pareto-improve the consumption bundle received by both workers and non-workers, a lottery market can be included to provide insurance against unemployment (i.e., not being selected for work) in a certain period. Such an arrangement would achieve full insurance (efficient risk sharing), so everyone would receive the same consumption independent of the employment status. If we assume that all households pool their resources together and maximize aggregate welfare, the resulting discounted utility function becomes as in Danthine and Kurmann (2004):

$$\sum_{t=0}^{\infty} \beta^t \{ \ln c_t - n_t G(e_t) \} \tag{2.3}$$

Each household starts with $k_0^i = k_0$ initial capital, which is equal to the aggregate capital in period 0. The aggregate capital stock then evolves as follows:

$$k_{t+1} = i_t + (1 - \delta)k_t \quad (2.4)$$

where $0 < \delta < 1$ denotes the depreciation rate on capital, and i_t is investment in period t . The before-tax rental rate on capital is r_t , and in addition the households have a legal claim on all the firm's profit π_t .

In addition to labor income, households receive labor income as well. The hourly wage rate in the economy is w_t , so the total before-tax labor income generated in each period is $w_t n_t$. The aggregate household's budget constraint is then

$$(1 + \tau^c)c_t + k_{t+1} - (1 - \delta)k_t = (1 - \tau^y)[r_t k_t + w_t n_t + \pi_t] + g_t^t \quad (2.5)$$

where τ^c is the consumption tax rate, τ^y is the common income tax rate, and g_t^t are the aggregate government transfers. The problem now is to maximize (2.3) s.t. (2.5). The first-order optimality conditions are as follows:

$$c_t: \frac{1}{c_t} = \Lambda_t(1 + \tau^c) \quad (2.6)$$

$$e_t: e_t = \phi_0 + (\phi_1 + \phi_3) \ln w_t + \phi_2 \ln n_t + \phi_4 \ln w_{t-1} \quad (2.7)$$

$$k_{t+1}: \Lambda_t = \beta E_t \Lambda_{t+1} [1 - \delta + (1 - \tau^y)r_{t+1}] \quad (2.8)$$

$$TVC: \lim_{t \rightarrow \infty} \beta^t \Lambda_t k_{t+1} = 0 \quad (2.9)$$

where Λ_t is the Lagrangian multiplier attached to the household's budget constraint in period t . The first optimality condition equates the marginal utility of consumption to the marginal utility of wealth. The second condition is called the Effort Condition (EC), or Solow (1979) condition. (Note that we have imposed the fact that in this economy everyone is identical, so the individual and the average wages are the same.) In other words, the optimal effort level depends on the changes in current and past wage rate, and current employment. The third condition is the employment optimality condition. The fourth is the so called Euler equation, which describes the optimal allocation of capital in any two congruent periods. The last condition, the Transversality condition (TVC), is a boundary condition that needs to be imposed to eliminate explosive solutions.

Firm

There is a stand-in firm that produces a homogeneous final good that can be used for consumption, investment, or government purchases. The Cobb-Douglas production function uses physical capital and efficiency labor as inputs as follows:

$$y_t = A_t k_t^{1-\alpha} (e_t n_t)^\alpha \tag{2.10}$$

where A_t captures the level of technology, $0 < \alpha, 1 - \alpha < 1$ are the efficiency labor and capital shares, respectively.

The firm maximizes profit subject to the household’s participation condition and effort condition being satisfied, which turns the firm’s problem becomes dynamic. More specifically, this is because the wage set today influences effort next period through the existence of past wage, w_{t-1} as an argument in the effort condition. The firm discounts profit by the stochastic discount factor (expressed in utility terms) $\Lambda_t = 1/[(1 + \tau^c)c_t]$, hence the firm’s dynamic problem is as follows:

$$\max \sum_{t=0}^{\infty} \beta^t \Lambda_t \{y_t - w_t n_t - r_t k_t\} \tag{2.11}$$

The resulting first-order conditions are

$$k_t: (1 - \alpha) \frac{y_t}{k_t} = r_t \tag{2.12}$$

$$n_t: \alpha \frac{y_t}{n_t} + \alpha \frac{y_t}{e_t} \frac{\partial e_t}{\partial n_t} = w_t \tag{2.13}$$

$$w_t: \alpha \frac{y_t}{e_t} + \beta \frac{\Lambda_{t+1}}{\Lambda_t} \alpha \frac{y_{t+1}}{e_{t+1}} \frac{\partial e_{t+1}}{\partial w_t} \tag{2.14}$$

where (2.13) and (2.14) make use of the fact that effort responds to the current and past wage rates. The first condition describes optimal renting of capital: in equilibrium it receives its marginal product. The second condition characterizes labor demand by the firm: in this setup, there is an elasticity term, $(\frac{\partial e}{\partial n})(\frac{n}{e}) \geq 0$, which appears to capture the effect of a new margin of adjustment. More specifically, a higher level of employment, though costly in terms of labor productivity, may actually increase the value of the firm’s investment and in turn worker’s effort; Using an analogy from finance, from the firm’s point of view, the worker is a multi-period asset. In other words, given the dynamic implications of the wage on the effort level exerted, the firm is hiring more people as compared to the perfectly competitive, perfect effort observability case.

The last equation describes how efficiency wages are set, i.e., how the firm chooses a wage rate to inspire the worker to supply optimum effort. Combining the optimality condition for employment and wages produces as in Vasilev (2017c):

$$1 = \epsilon(e_t, w_t) - \epsilon(e_t, n_t) + \beta E_t \left[\frac{\Lambda_{t+1}}{\Lambda_t} \frac{y_{t+1}}{y_t} \epsilon(e_{t+1}, w_t) \right] \tag{2.15}$$

Where $\epsilon(e_t, w_t)$ denotes the elasticity of effort level with respect to the wage rate, and $\epsilon(e_{t+1}, w_t)$ denotes the elasticity of the next-period effort level with respect to the current wage. With $\epsilon(e_t, w_t) > 0$, the standard Solow (1979) condition does not apply, since an increase in the wage rate at the margin produces an extra increase in worker's productivity; similarly, with $\epsilon(e_{t+1}, w_t) < 0$, the firm has to take into consideration the future effect of the current wage rate - that a higher wage paid today makes it more costly to extract higher effort from a worker in the future.

Given the assumed functional form, we can now solve for the optimal effort level to obtain

$$e_t = \phi_1 + \phi_3 - \phi_2 + \beta\phi_4 = \bar{e} > 0 \quad (2.16)$$

That effort is positive and constant over the cycle. Next, plugging that expression into the effort function yields

$$\ln w_t = \gamma_0 + \gamma_1 \ln n_t + \gamma_2 \ln w_{t-1} \quad (2.17)$$

where

$$\gamma_0 = \frac{\phi_1 + \phi_3 - \phi_2 + \beta\phi_4 - \phi_0}{\phi_1 + \phi_3}, \gamma_1 = -\frac{\phi_2}{\phi_1 + \phi_3}, \gamma_2 = -\frac{\phi_4}{\phi_1 + \phi_3} \quad (2.18)$$

which differs from Danthine and Kurmann (2004); Here we are following Danthine and Kurmann (2010), and internalize the effect that past wages have on current effort level, which is the more realistic case.

Government

The government will be assumed to be running a balanced budget in every period. The government collects revenue from levying taxes on capital, labor, and profit income, and then spends on government consumption and transfers, which are returned lump-sum to the households:

$$\tau^c c_t + \tau^y [r_t k_t + w_t n_t + \pi_t] = g_t^c + g_t^t \quad (2.19)$$

Where g_t^c are (non-productive) government purchases. Government spending share will be set equal to its long-run average so that the level will be varying with output. Government transfers will be residually determined and will always adjust to make sure that the government budget is balanced.

Stochastic Processes for the exogenous variables

The first exogenous stochastic variable is total factor productivity A_t , which is assumed to follow AR(1) processes in logs, in particular

$$\ln A_{t+1} = (1 - \rho_a) \ln A_0 + \rho_a \ln A_t + \varepsilon_{t+1}^a \quad (2.20)$$

where $A = A_0 > 0$ is the steady-state level of total factor productivity process, $0 < \rho_a < 1$ is the first-order autoregressive persistence parameter and $\varepsilon_{t+1}^a \sim iidN(0, \sigma^2)$ are ran-

dom shocks to the total factor productivity process. Hence, the innovations to technology represent unexpected changes in the total factor productivity process.

Decentralized Dynamic Equilibrium with fair wages

Given the process followed by total factor productivity $\{A_t\}_{t=0}^{\infty}$, the tax rates $\{\tau^c, \tau^y\}$, initial capital stock k_0 , hours worked per household h^i , the decentralized dynamic equilibrium with fair wages is a list of sequences $\{c_t, l_t, k_t, e_t\}_{t=0}^{\infty}$ for each household i , input levels $\{k_t, n_t, e_t\}_{t=0}^{\infty}$ chosen by the firm, a sequence of government spending categories $\{g_t^c, g_t^t\}_{t=0}^{\infty}$ and input prices $\{w_t, r_t\}_{t=0}^{\infty}$ such that (i) each household i maximizes its utility function subject to its budget constraint; (ii) the representative firm maximizes profit by offering a wage schedule that satisfies the workers' incentive compatibility constraint and to induce an optimal effort level; (iii) government budget is balanced in each period; (iv) all markets clear.

Data and model calibration

To characterize business cycle fluctuations in Bulgaria, we will focus on the period following the introduction of the currency board (1999-2016). Quarterly data on output, consumption, and investment was collected from the National Statistical Institute (2017), while the real interest rate is taken from Bulgarian National Bank Statistical Database (2017). The calibration strategy described in this section follows a long-established tradition in modern macroeconomics: first, as in Vasilev (2016), the value of the discount factor, $\beta = 0.982$, is set to match the steady-state capital-to-output ratio in Bulgaria, $k/y = 13.964$, in the steady-state Euler equation. The labor share parameter, $\alpha = 0.571$, is obtained as in Vasilev (2017d), and equals the average value of labor income in aggregate output over the period 1999-2016. This value is slightly higher as compared to other studies on developed economies, due to the overaccumulation of physical capital, which was part of the ideology of the totalitarian regime, which was in place until 1989. Next, the average income tax rate was set to $\tau^y = 0.1$. This is the average effective tax rate on income between 1999-2007 when Bulgaria used progressive income taxation, and equal to the proportional income tax rate introduced as of 2008. Similarly, the tax rate on consumption is set to its value over the period, $\tau^c = 0.2$. This is in line with the estimates for Bulgaria (Vasilev 2017a) as well over the period studied.

Next, the steady-state depreciation rate of physical capital in Bulgaria, $\delta = 0.013$, was taken from Vasilev (2016). It was estimated as the average quarterly depreciation rate over the period 1999-2014. Next, the steady state employment rate in Bulgaria is set to $n = 0.533$, as in Vasilev (2016a). Following Vasilev (2017c), the values for γ_1 and γ_2 were estimated from the optimized effort function. Finally, the processes followed by total factor productivity (TFP) is estimated from the detrended series by running an AR(1) regression and saving the residuals. Table 1 below summarizes the values of all model parameters used in the paper.

Table 1. Model Parameters

Param	Value	Description	Method
β	0.982	Discount factor	Calibrated
α	0.571	Labor Share	Data avg.
δ	0.013	Depreciation rate on physical capital	Data avg.
n	0.533	Employment rate	Data avg
γ_1	0.140	Weight on labor market conditions	Estimated
γ_2	0.814	Weight attached to past wage	Estimated
τ_y	0.100	Average tax rate on income	Data avg.
τ_c	0.200	VAT/consumption tax rate	Data avg.
ρ_a	0.701	AR(1) persistence coefficient, TFP	Estimated
σ_a	0.044	st. error, TFP process	Estimated

Source: Author's calculations

Steady-State

Once the values of model parameters were obtained, the steady-state equilibrium system solved, the "big ratios" can be compared to their averages in Bulgarian data. The results are reported in Table 2 on the next page. The steady-state level of output was normalized to unity (hence the level of technology A differs from unity), which greatly simplified the computations, and allows the steady-state to be solved by hand. Next, the model matches consumption-to-output ratio by construction; The investment and government purchases ratios are also closely approximated. The shares of income are also identical to those in data, which follows directly from the constant-returns to scale featured by the aggregate production function. The after-tax return, where $\tilde{r} = (1 - \tau_y) - \delta$ is also relatively well- captured by the model.

Table 2. Data Averages and Long-run Solution

Variable	Description	Data	Model
y	Steady-state output	N/A	1.000
c/y	Consumption-to-output ratio	0.648	0.674
i/y	Investment-to-output ratio	0.201	0.175
k/y	Capital-to-output ratio	13.96	13.96
gc/y	Government consumption-to-output ratio	0.151	0.151
wh/y	Labor income-to-output ratio	0.571	0.571
rk/y	Capital income-to-output ratio	0.429	0.429
h	Share of time spent working	0.333	0.333
rb	After-tax net return on capital/bond rate	0.014	0.016

Source: Author's calculations.

Out of steady-state model dynamics

Since the model does not have an analytical solution for the equilibrium behavior of variables outside their steady-state values, we need to solve the model numerically. This is done by log-linearizing the original equilibrium (non-linear) system of equations around the steady-state. This transformation produces a first-order system of stochastic difference equations. First, we study the dynamic behavior of model variables to an isolated shock to the total factor productivity process, and then we fully simulate the model to compare how the second moments of the model perform when compared against their empirical counterparts.

Impulse Response Analysis

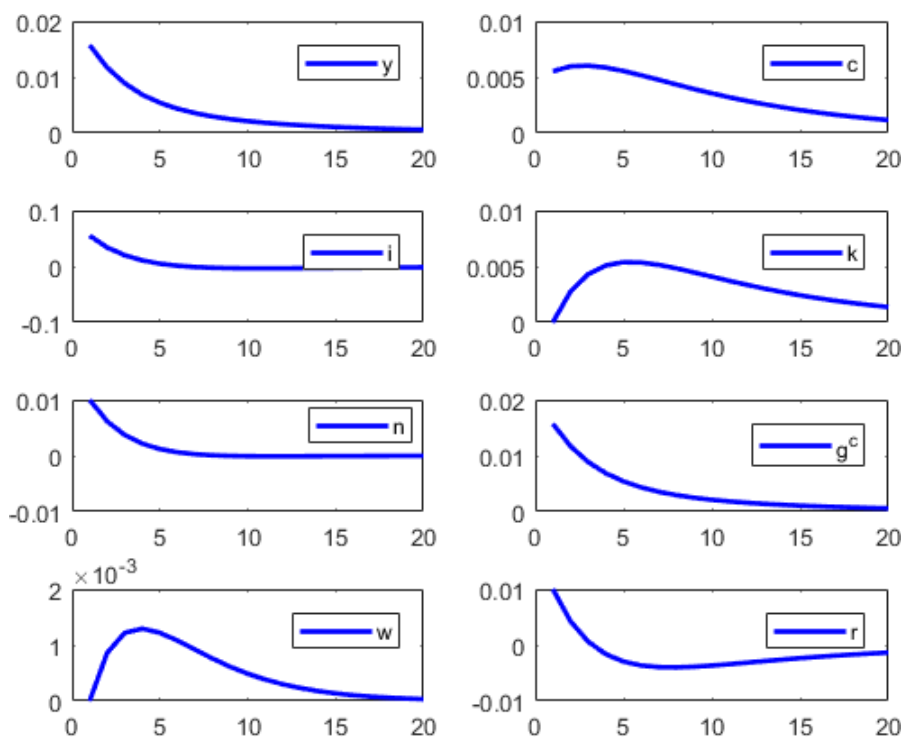
This subsection documents the impulse responses of model variables to a 1% surprise innovation to technology. The impulse response functions (IRFs) are presented in Fig. 1 on the next page. As a result of the one-time unexpected positive shock to total factor productivity, output increases upon impact. This expands the availability of resources in the economy, so use of output - consumption, investment, and government consumption also increase contemporaneously.

At the same time, the increase in productivity increases the after-tax return on the two factors of production, labor, and capital. The representative households then respond to the incentives contained in prices and start accumulating capital, and supplies more hours worked. In turn, the increase in capital input feeds back in output through the production function, and that further adds to the positive effect of the technology shock. In the labor market, the wage rate increases, and the household increases its hours worked. In turn, the increase in total hours further increases output, again indirectly.

Over time, as capital is being accumulated, its after-tax marginal product starts to decrease, which lowers the households' incentives to save. As a result, the physical capital stock eventually returns to its steady-state and exhibits a hump-shaped dynamics over its transition path. The rest of the model variables return to their old steady-states in a monotone fashion as the effect of the one-time surprise innovation in technology dies out.

Simulation and moment matching

As in Vasilev (2017b), we will now simulate the model 10,000 times for the length of the data horizon. Both empirical and model simulated data is detrended using the Hodrick-Prescott (1980) filter. Table 3 on the next page summarizes the second moments of data (relative volatilities to output, and contemporaneous correlations with output) versus the same moments computed from the model-simulated data at a quarterly frequency. To minimize the sampling error, the simulated moments are averaged out over the computer-generated draws. As in Vasilev (2016, 2017b, 2017c), the model matches quite well the absolute volatility of output. By construction, government consumption in the model varies as much as output. The model overestimates the variability in consumption, and less so, the predicted volatility of the investment. Still, the model is qualitatively consistent with the stylized fact that consumption generally varies less than output, while investment is more volatile than output.

Figure 1. Impulse Responses to a 1% surprise innovation in technology

Source: Author's calculation

With respect to the labor market variables, the variability of employment, unemployment, and wages predicted by the model is too low than that in data. This is a confirmation that the incentive wage function, when calibrated to Bulgarian data, does not describe very well the dynamics of labor market variables after the introduction of the currency board. Next, in terms of contemporaneous correlations, the model systematically over-predicts the procyclicality of the main aggregate variables - investment, and government consumption. This, however, is a common limitation of this class of models. With respect to wages, the model predicts strong cyclicity, while wages in data are acyclical. This shortcoming is well-known in the literature and an artifact of the wage following closely labor productivity in the model. In the next subsection, as in Vasilev (2016), we investigate the dynamic correlation between labor market variables at different leads and lags, thus evaluating how well the model matches the phase dynamics among variables. In addition, the autocorrelation functions (ACFs) of empirical data, obtained from an unrestricted VAR(1) are put under scrutiny and compared and contrasted to the simulated counterparts generated from the model.

Table 3. Business Cycle Moments

	Data	Model
σ_y	0.05	0.05
σ_c/σ_y	0.55	0.90
σ_i/σ_y	1.77	1.92
σ_g/σ_y	1.21	1.00
σ_h/σ_y	0.63	0.37
σ_w/σ_y	0.83	0.14
$\sigma_y/h/\sigma_y$	0.86	0.14
σ_u/σ_y	3.22	0.37
$corr(c, y)$	0.85	0.93
$corr(i, y)$	0.61	0.80
$corr(g, y)$	0.31	1.00
$corr(h, y)$	0.49	0.93
$corr(w, y)$	-0.01	0.96
$corr(u, y)$	-0.47	-0.82

Source: Author's calculation

Auto- and cross-correlations

This subsection discusses the auto-(ACFs) and cross-correlation functions (CCFs) of the major model variables. The coefficients' empirical ACFs and CCFs at different leads and lags are presented in Table 4 below against the averaged simulated AFCs and CCFs. Following Canova (2007), this is used as a goodness-of-fit measure.

Table 4: Autocorrelations for Bulgarian data and the model economy

		k			
Method	Statistic	0	1	2	3
Data	$corr(u_t, u_{t-k})$	1.000	0.765	0.552	0.553
Model	$corr(u_t, u_{t-k})$	1.000	0.954	0.898	0.833
	(s.e.)	(0.000)	(0.028)	(0.055)	(0.080)

Data	$corr(n_t, n_{t-k})$	1.000	0.484	0.009	0.352
Model	$corr(n_t, n_{t-k})$	1.000	0.954	0.898	0.833
	(s.e.)	(0.000)	(0.028)	(0.055)	(0.080)
Data	$corr(y_t, y_{t-k})$	1.000	0.810	0.663	0.479
Model	$corr(y_t, y_{t-k})$	1.000	0.957	0.906	0.847
	(s.e.)	(0.000)	(0.027)	(0.051)	(0.075)
Data	$corr(a_t, a_{t-k})$	1.000	0.702	0.449	0.277
Model	$corr(a_t, a_{t-k})$	1.000	0.956	0.903	0.840
	(s.e.)	(0.000)	(0.027)	(0.053)	(0.077)
Data	$corr(c_t, c_{t-k})$	1.000	0.971	0.952	0.913
Model	$corr(c_t, c_{t-k})$	1.000	0.958	0.909	0.853
	(s.e.)	(0.000)	(0.026)	(0.049)	(0.072)
Data	$corr(i_t, i_{t-k})$	1.000	0.810	0.722	0.594
Model	$corr(i_t, i_{t-k})$	1.000	0.953	0.894	0.825
	(s.e.)	(0.000)	(0.029)	(0.056)	(0.082)
Data	$corr(w_t, w_{t-k})$	1.000	0.760	0.783	0.554
Model	$corr(w_t, w_{t-k})$	1.000	0.958	0.906	0.846
	(s.e.)	(0.000)	(0.027)	(0.051)	(0.075)

Source: Author's calculation

As seen from Table 4 above, the model compares relatively well vis-a-vis data. Empirical ACFs for output and investment are slightly outside the confidence band predicted by the model, while the ACFs for total factor productivity and household consumption are well- approximated by the model. The persistence of labor market variables is also relatively well-described by the model dynamics. Overall, the model with habits in consumption generatess too much persistence in output and both employment and unemployment, and is subject to the criticism in Nelson and Plosser (1992), Cogley and Nason (1995) and Rotemberg and Woodford (1996b), who argue that the RBC class of models do not have a strong internal propagation mechanism besides the strong persistence in the TFP process. In those modelss, e.g., Vasilev (2009), and in the current one, labor market is modelled in the Walrasian market-clearing spirit, and output and unemployment persistence are low.

Next, as seen from Table 5 below, over the business cycle, in data labor productivity leads employment. The model, however, cannot account for this fact, despite the dependence of the wage function on its lagged value. As in the standard RBC model a technology shock can be regarded as a factor shifting the labor demand curve, while holding the labor supply curve constant. Therefore, the effect between employment and labor productivity is only a contemporaneous one.

Table 5. Dynamic correlations for Bulgarian data and the model economy

		k						
Method	Statistic	-3	-2	-1	0	1	2	3
Data	$corr(n_t, (y/n)_{t-k})$	-0.342	-0.363	-0.187	-0.144	0.475	0.470	0.346
Model	$corr(n_t, (y/n)_{t-k})$	0.081	0.104	0.133	0.819	0.288	0.183	0.106
	(s.e.)	(0.314)	(0.273)	(0.227)	(0.149)	(0.240)	(0.272)	(0.303)
Data	$corr(n_t, w_{t-k})$	0.355	0.452	0.447	0.328	-0.040	-0.390	-0.57
Model	$corr(n_t, w_{t-k})$	0.081	0.104	0.133	0.819	0.288	0.183	0.106
	(s.e.)	(0.314)	(0.273)	(0.227)	(0.149)	(0.240)	(0.272)	(0.303)

Conclusions

In this paper, we investigate the quantitative importance of "fair" wages in explaining fluctuations in Bulgarian labor markets. In contrast to Vasilev (2017b) who introduces wages of no-shirking type a la Shapiro and Stiglitz (1984), and Vasilev (2017c), who incorporated reciprocity arrangement in labor relations and wage determination a la "gift exchange" as in Akerloff (1982), into a relatively standard RBC model with government sector, here we follow Danthine and Kurmann (2004, 2010). We calibrated the model to Bulgarian data after the introduction of the currency board arrangement and studied the impulse responses of aggregate variables in the face of exogenous technological shocks. Overall, the calibrated model with "fair" wages performs poorly when it comes to the relative volatilities of labor market variables. This is because aggregate labor market conditions, as proxied by the employment rate in the economy, turn out to be of lesser importance. As suggested in Vasilev (2017c), rent-sharing aspects, and firms' ability to pay, are the most important determinants of wages in Bulgaria over the period 1999-2015. Similar findings have been documented in Lozev, Vladova and Paskaleva (2011) and Paskaleva (2016).

Disclosure statement: No potential conflict of interest was reported by the author.

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