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Exchange rates and prices: A continuous wavelet perspective

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MT-DP - 2018/33

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GÁBOR ULIHA - JÁNOS VINCZE

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Exchange rates and prices: a continuous wavelet perspective

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Exchange rates and prices: a continuous wavelet perspective

Gábor Uliha - János Vincze

Abstract

In this paper we analyze statistics derived from the cross-wavelet transform of inflation differentials and exchange rate changes for a group of countries with Germany as the reference country. An important tool is the wavelet coherency measure from which we can judge the strength of the price-exchange rate nexus at different time scales, and also whether it has changed in time. Complex cross-wavelets provide information about phase relationship, and we can investigate whether there is any consistent pattern in the lead-lag relationship between prices and exchange rates. Also, we calculate a summary measure, based on singular value decomposition, that shows which countries have significantly similar inflation differential – exchange rate change processes. Our results accord, in some ways, with former findings, but suggest an even gloomier view on the possibility of finding statistically reliable relationships between exchange rates and aggregate price indices (CPI or PPI). In line with the literature we haven't found strong co-movement between prices and exchange rates in the short or medium term. There are only weak indications that at least in some countries the price-exchange rate connection strengthened during the crisis, and detectable cycles at business cycle frequencies do not appear at all. Though by and large the lead-lag relationship between prices and exchange rates is the expected one, still this is unstable practically for every country. Results with respect to the PPI are more promising. The three countries that seem to be closest to theoretical expectations are Sweden, Japan and South-Korea. It is possible that the coherence between exchange rates and prices on the macro level may depend more on similarities of export structure than on trading relations, that microeconomic intuition would suggest. Also it is possible that macro price indices are too noisy by their very nature to be amenable for statistical analyses not committed to strong presumptions.

Keywords: real exchange rates, exchange rate pass-through, continuous wavelet analysis

JEL-codes: E31, C14, F41

Árfolyamok és árak: egy folytonos wavelet elemzés

Uliha Gábor - Vincze János

Összefoglaló

Ebben a tanulmányban a nominális árfolyamok és az inflációs különbségek összefüggését

elemezzük 19 országra a folytonos wavelet transzformáció segítségével. Referencia országnak

Németországot választottuk. Legfőbb analitikus eszközünk a wavelet koherencia, amely képes

a kapcsolatok erősségét időtávokra és időpontokra is értelmezni, továbbá számot adni arról,

hogy melyik idősor vezet és melyik követ. Különböző összegző mértékeket is kiszámítottunk,

amelyekből a kapcsolatok általános erősségére tehetünk megállapításokat. Eredményeink sok

tekintetben hasonlóak a más módszerekkel elért kvalitatív megállapításokhoz, de bizonyos

értelemben sötétebb képet festenek, amennyiben kevés konzisztens minta létszik

kirajzolódni, még olyan esetekben is, amelyekben ilyet a priori elvárnánk. (Ilyen például

Dánia esete.) Sem rövid, sem középtávon nem találunk általában szoros kapcsolatot, és nem

tűnik úgy, hogy a válság felerősítette volna az együttmozgást. Bár a vezető-követő kapcsolat

nagyjából megfelel a várakozásoknak, túlságos instabil ahhoz, hogy megbízható előrejelzések

alapjául szolgálhatna. A CPI-hez képest a PPI-vel elért eredmények inkább bíztatók, ismét

összhangban az eddigi irodalommal. Svédország, Japán és Dél-Korea az a három ország,

amely leginkább az elméleti várakozásoknak megfelelően viselkedik. Ez arra utalhat, hogy

nem annyira a kereskedelem intenzitása, mint inkább strukturális hasonlóságok játszhatnak

szerepet az aggregált árindexek és az árfolyamok együttmozgásában.

Tárgyszavak: reálárfolyam, árfolyam-ár kapcsolat, folytonos wavelet elemzés

JEL kódok: E31, C14, F41

Exchange rates and prices: a continuous wavelet perspective

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Abstract

In this paper we analyze statistics derived from the cross-wavelet transform of inflation differentials and exchange rate changes for a group of countries with Germany as the reference country. An important tool is the wavelet coherency measure from which we can judge the strength of the price-exchange rate nexus at different time scales, and also whether it has changed in time. Complex cross-wavelets provide information about phase relationship, and we can investigate whether there is any consistent pattern in the lead-lag relationship between prices and exchange rates. Also, we calculate a summary measure, based on singular value decomposition, that shows which countries have significantly similar inflation differential – exchange rate change processes. Our results accord, in some ways, with former findings, but suggest an even gloomier view on the possibility of finding statistically reliable relationships between exchange rates and aggregate price indices (CPI or PPI). In line with the literature we haven't found strong co-movement between prices and exchange rates in the short or medium term. There are only weak indications that at least in some countries the price-exchange rate connection strengthened during the crisis, and detectable cycles at business cycle frequencies do not appear at all. Though by and large the lead-lag relationship between prices and exchange rates is the expected one, still this is unstable practically for every country. Results with respect to the PPI are more promising. The three countries that seem to be closest to theoretical expectations are Sweden, Japan and South-Korea. It is possible that the coherence between exchange rates and prices on the macro level may depend more on similarities of export structure than on trading relations, that microeconomic intuition would suggest. Also it is possible that macro price indices are too noisy by their very nature to be amenable for statistical analyses not committed to strong presumptions.

JEL: E31, C14, F41

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1 Introduction

The real exchange rate is an important macroeconomic variable. It is rarely regarded as an innocuous relative price, just the equilibrium value of a unit of country A's goods in terms of Country B's goods. Rather it is considered as having a long-run equilibrium value from which it may deviate, and these deviations can signal some underlying international imbalance. Also it is thought by many economists that the real exchange rate can be easily influenced by policy, thereby it can be used as a means to achieve certain goals. The real exchange rate can be factored as the nominal exchange rate times the ratio of price levels measured in domestic prices, thus its percentage change can be written as the sum of currency revaluation and the inflation differential between the countries in question.

This factorization raises questions about the relationship between pricing and the nominal exchange rate. Frequently the question is formulated as that of the exchange rate pass-through, i.e. how fully and at what horizons exchange rate changes translate into changes in international relative prices. In, practically all theories of international economics there exists some relationship between exchange rates and international prices, but common sense would also tell us the same. At the microeconomic level connections must exist, but these may be variable across markets, countries, and time, depending on transaction costs, the nature of competition etc. Therefore, one could not expect that aggregate statistics exhibit any stable relationship, and could make justice to any theory of international price determination, still it has repeatedly been attempted.

Even if we cannot expect to find some deep, structural relationship the statistical exploration of exchange rate changes and inflation differentials is still worthwhile. The relationship is an important one for forecasters, and for policy exercises, even without having access to the underlying processes. Certainly, for the practical use of the monetary transmission mechanism it is imperative to know at what horizons exchange rates and prices are decoupled.

In this paper our goal is to document the statistical properties of the "exchange-rate pass-through" in a framework where its stability in time is not a maintained hypothesis. In fact, we look for evidence whether some of the qualitative features that have been established are valid (in a qualitative sense) from the point of view of a broader minded maintained hypothesis than the ones usually employed in the past. More concretely most statistical analysis have been conducted on the assumption that there exist equilibrium real exchange rates. This does not imply that relative prices and nominal exchange rates are co-integrated, as the latter assumption is in general discredited at the moment. It is a more popular assumption to hypothesize a time-varying equilibrium real exchange rate, depending on factors like differential productivity growth (see Égert et al. (2006), Bordo et al. (2017)). The usual results in the literature include (see for instance Burstein-Gopinath (2014), Taylor-Taylor (2004)): in the short-term inflation and exchange rates are decoupled, but in the long term pass-through is significantly positive (going from exchange rates to prices). Half-lives are measured at the scale of several years, but there are differences with respect to price aggregates, for instance, producer prices adjust more quickly and strongly than consumer prices.

In this paper we want to address these issues with the help of a methodology that can handle non-linearities, non-stationarities, and is able to separate relationships on different horizons. Our analysis is based on the continuous cross-wavelet transform, a statistical tool gaining popularity in the time-series literature for describing dependencies between time-series. It allows us to attack the time scale problem of the dependency, and whether there were structural changes in the relationship, as well as the very existence of an association. It can tell us something also about the lead-lag relationship. As wavelets can be regarded as an alternative vantage point, there is no exact correspondence between the traditional findings and ours, we have to avail ourselves by making qualitative judgments.

In the next section we overview continuous wavelet theory, emphasizing those aspects that we will make use of in the analysis. Section 3 describes our data, and the results of the analysis, while Section 4 summarizes.

2 Why wavelets?

As a simplification one could say that the wavelet transform expresses how much a time series changed around a certain date at different scales (frequencies). It has been likened to a prism through which one can observe the properties of an object (the time series in our case) otherwise obscured. It is customary to relate it to the Fourier-transform that assumes a similar task, but relies on the assumption of homogeneity (stationarity), and does not account for local (localized in time) changes. In the role of prism wavelets have been proved to improve on Fourier-analysis, at least in the life and earth sciences. In other words, to characterize complex and non-stationary systems this methodology has advantages. (See Torrence-Compo (1998).)

Let x(t) be square integrable then its continuous wavelet transform at time τ , and scale s is:

$$W_{x}(\tau,s) = \int_{-\infty}^{\infty} x(t)\psi_{\tau,s}^{*}(t)dt$$
 /1/

Here $\psi_{\tau,s}(t)$ denotes the wavelets, and * the complex conjugate. Besides $s, T \in \mathbb{R}, s \neq 0$. The the wavelets are derived from the the mother wavelet $(\psi(\cdot))$ as:

$$\psi_{\tau,s}(t) = s^{-0.5}\psi\left(\frac{t-\tau}{s}\right)$$
 /2/

For a mother wavelet $\psi(\cdot) \in L^2(\mathbb{R})$ is a condition, and also for reconstructibility

$$0 < \int_{-\infty}^{\infty} \frac{|\Psi(\omega)|}{|\omega|} d\omega < \infty$$
 /3/

is usually assumed.

Continuous wavelets are highly redundant transformations, when calculated from an actual time series the computation produces a matrix with much more entries than the original series. They must be distinguished from discrete wavelets that specifically strive for data compression and ate used much less in research than in engineering.

In economic applications the most commonly used mother wavelet is the Morlet wavelet (Goupillaud et al. [1984])¹:

$$\psi(t) = \frac{\pi^{-0.25}e^{-0.5t^2}}{\cos(\omega t) - i \cdot \sin(\omega t)}$$
/5/

For the reasons using the Morlet-wavelet see (),(), here we take notice of two aspects: 1. with the choice of ω =6 one can interpret scale as frequency, and 2. as it is a complex wavelet we can have phase difference information when applying the methodology to two different series.

The definition of the cross-wavelet transform (*Hudgins et al.* [1993]):

$$W_{x,y}(\tau,s) = W_x(\tau,s)W_y^*(\tau,s)$$
 /4/

¹ This formula does not satisfy the recoverability criterion, a correction term is needed for that. (*Foufoula-Georgiou–Kumar* [1994]).

The cross-wavelet transform serves to calculate local "covariances" over different time scales. The wavelet coherence is a non-signed expression, similar to (local) correlation.

$$R_{x,y}(\tau,s) = \frac{\left| S\left(W_{x,y}(\tau,s) \right) \right|}{\left\{ S[|W_x(\tau,s)|^2] S\left[|W_y(\tau,s)|^2 \right] \right\}^{0.5}}$$
 /5/

Unfortunately, to calculate it we need a smoothing function S(.), otherwise the coherency measure would always be 1^2 .

In addition, one can measure the lead-lag relationship between two series with the help of the phase difference:

$$\phi_{x,y}(\tau,s) = \arctan\left(\frac{\Im\left(S\left(W_{x,y}(\tau,s)\right)\right)}{\Re\left(S\left(W_{x,y}(\tau,s)\right)\right)}\right)$$
 /6/

Here $\Im(\cdot)$ denotes the imaginary, and $\Re(\cdot)$ the real part. The phase difference can take values in the interval $[-\pi;\pi]$ and indicates the following lead-lag relationships (see, for instance, Aguilar-Conraria-Soares (2011).):

$$\phi_{x,y}(\tau,s) \in \left(-\frac{\pi}{2};0\right)$$
: the series are *in-phase*, and y leads

$$\phi_{x,y}(\tau,s) \in \left(0,\frac{\pi}{2}\right)$$
: the series are *in-phase*, and x leads

$$\phi_{x,y}(\tau,s) \in \left(\frac{\pi}{2};\pi\right)$$
: the series are *out-of-phase*, and y leads

$$\phi_{x,y}(\tau,s) \in \left(-\pi,-\frac{\pi}{2}\right)$$
: the series are *out-of-phase*, and x leads.

What kind of statistics can we derive that would characterize our inflation - exchange rate data with a view towards the questions posed in the Introduction? The Wavelet Power Spectrum is the squared wavelet transform. The WPS figures we are going to report have the following interpretation: a point with abscissa (time period), and ordinate (scale) expresses the power attributable to that time and scale. The coloring reflects the usual coloring of heat maps, deep red indicating strong impact and light blue no impact. The integral of the WPS equals the variance of the time series, thus the figure can be interpretable as variance decomposition, as well.

The WC (Wavelet Coherency) figures show the coherency measures for a given time and scale with the same coloring convention. Here the value is between 0 and 1. On these figures one can notice areas with dark borders, these are those where the coherency is significant at the 5 % level. (About the null hypothesis and the test see ().) Also light lines indicate the borders of an area called the cone of influence. Outside the cone of influence edge effects make dubious the estimates. Arrows are attached to areas with 5 % significance. The direction of the arrows corresponds to phase differences, with the natural convention. (For instance, a south-easterly arrow corresponds to a radian between -pi/2 and 0, in other words: x and y are in-phase and y leads.)

_

² We use the Hamming-window (*Harris* [1978]).

Beside these images we calculate the most powerful time and the most powerful scale statistics, where the WC values are averaged time and scale-wise, and then arg-maxed according to time and scale, respectively. We also report summary measures of the strength of the association by computing the percentage of the significant (at the 5 % level) area within each WC.

We report another one dimensional measure of association, as suggested in Soares-Aguira-Conraria (2014). To obtain this one has to compute the SVD (singular value decomposition) of the WC. The SVD provide leading patterns (l_x^k , l_y^k , u_x^k , u_y^k) for the constituent wavelets, of which they can be approximately reconstructed as

$$W_x \approx \sum_{k=1}^K u_x^k l_x^k \text{ és } W_y \approx \sum_{k=1}^K u_y^k l_y^k.$$

Next the similarity index is calculated as follows:

$$dist(W_{x}, W_{y}) = \frac{\sum_{k=1}^{K} \sigma_{k}^{2} [d(l_{x}^{k}, l_{y}^{k}) + d(u_{x}^{k}, u_{y}^{k})]}{\sum_{k=1}^{K} \sigma_{k}^{2}}.$$

3 Data and results

The source of our data is the International Financial Statistics. Exchange rate and inflation (CPI and PPI) data are monthly, covering the periods 1991-2017. We take Germany as the reference country, in other words exchange rates refer to DM and euro exchange rates of the countries in the sample, and inflation differences are defined with respect to the German PPI and CPI. We selected Germany as the reference because of its size in international trade, and because of its variation in geographical proximity to many other countries in the sample. Our sample contain countries belonging to the European Union, several large countries with an important share in world trade, and also less important countries internationally that, however, are not distant from Germany. The list includes in alphabetical order: Algeria, Canada, China, Denmark, Egypt, Hungary, Iceland, Israel, Japan, South Korea, Morocco, Norway, Poland, Sweden, Switzerland, Tunisia, Turkey, United Kingdom, United States.

Results

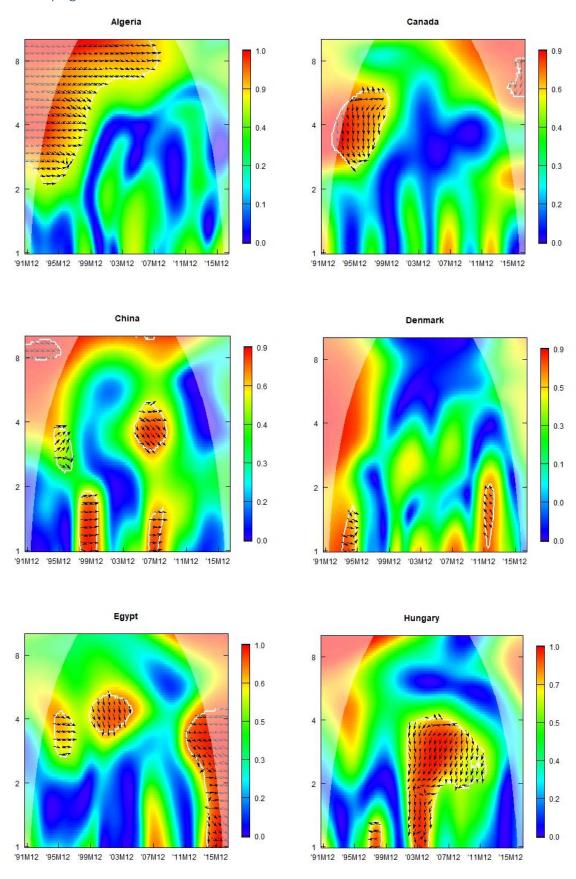
The following figures show coherencies for the 19 countries listed above. The first type shows the coherency between the CPI inflation differential vis a vis Germany, and the nominal euro (DM) rate log change. The second set exhibits the same statistics with the PPI instead of the CPI. As fewer PPI series are available there remain only 12 countries. In the parlance of the previous section the inflation differential is the variable x, and the log exchange rate change is y.

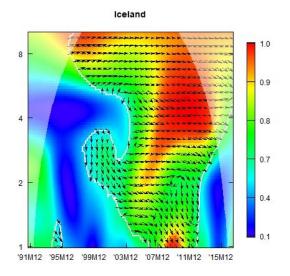
Phase differences

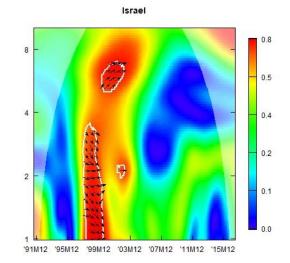
Our expectation is that exchange rates lead prices, and they are in-phase rather than out-of-phase. In visual terms arrows must be directed towards the South-East. Inspecting Figures ??? we can judge whether this expectation is fulfilled. On these figures the arrows are shown only for the significant areas.

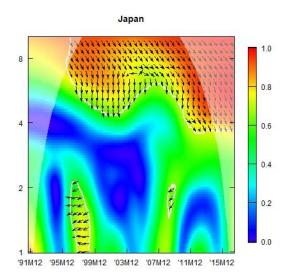
Though in general one would say that south-easterly is the dominant direction in no case can we find a totally consistent pattern. For several countries there are more or less substantial connected areas where the expected picture materializes. These countries include Algeria, Egypt, Iceland, Japan in the case of the CPI, and Hungary, Japan, Poland, South Korea, Sweden, Turkey and the UK with the PPI.

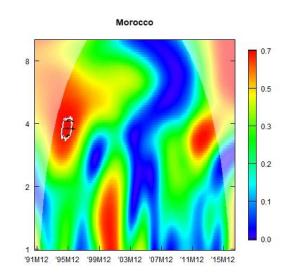
Coherency figures: CPI

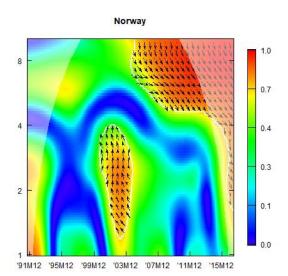


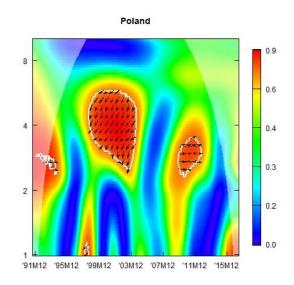


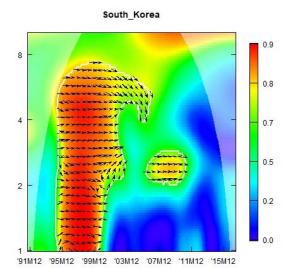


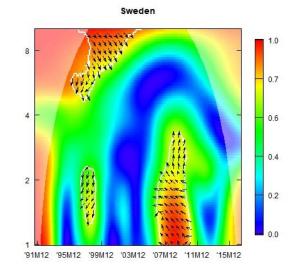


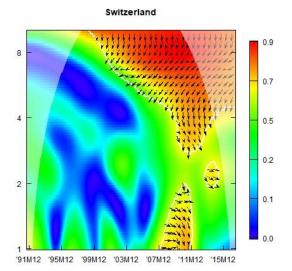


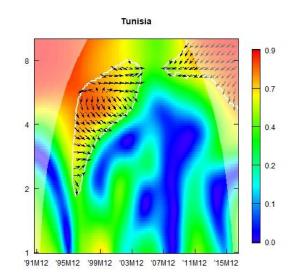


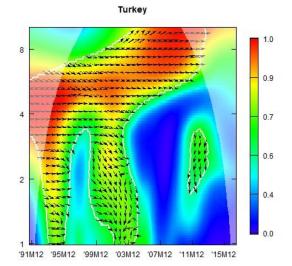


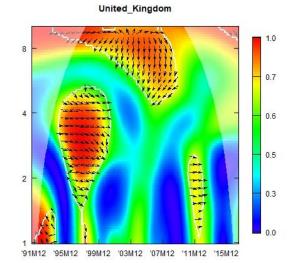


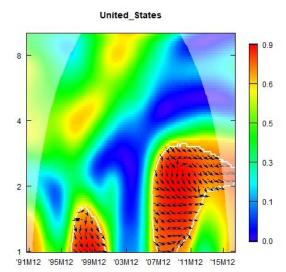


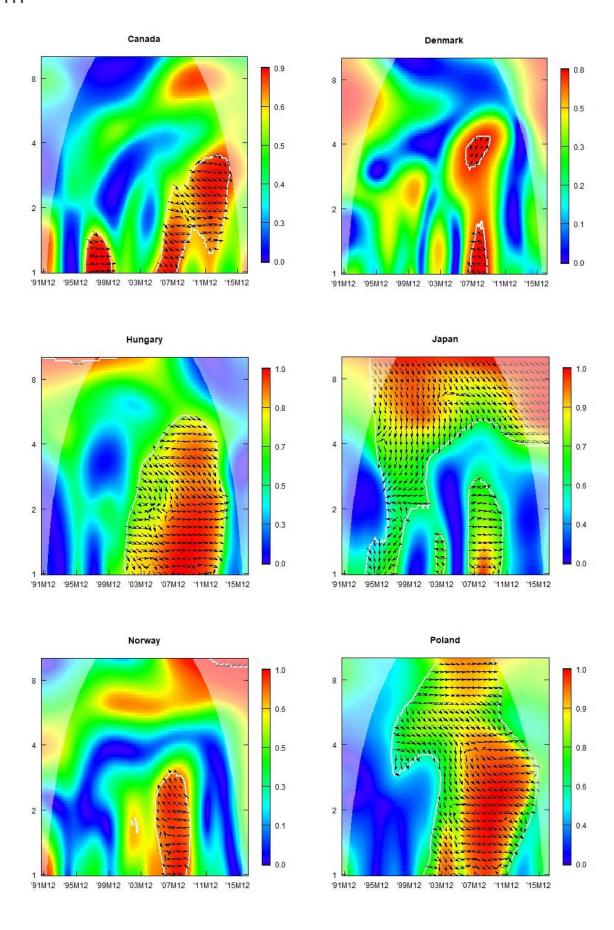


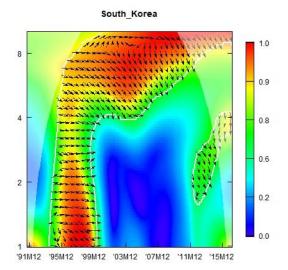


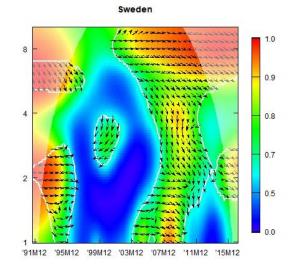


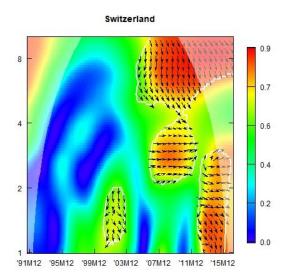


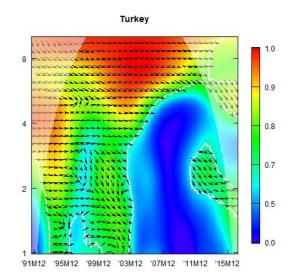


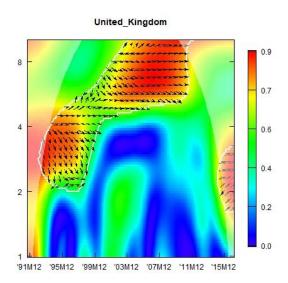


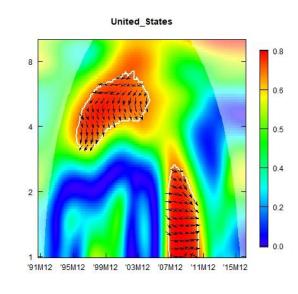












Percentage of significant area

The last three columns of Table 1 shows coherency significant area percentages at 1,5 and 10 %. This measure shows that the highest degree of coherence between exchange rates and the CPI prevails for Iceland and Turkey. Only in these two cases is the percentage above 50 % at the 100 % level. As Iceland is a small country closely related to the EU this does not seem to be surprising, but Turkey is not the country that one could expect to have the second place. The poorest performers are Morocco and Denmark, the latter should have been *a priori* in the class of Iceland rather, than in that of Morocco.

Table 1 Percentages of significant area: CPI

	Power - CPI inflation		Power - EXCH			Coherence (CPI-EXCH)			
	SIG01	SIG05	SIG10	SIG01	SIG05	SIG10	SIG01	SIG05	SIG10
Algeria	4,3%	11,4%	16,5%	7,6%	12,2%	17,3%	20,0%	25,3%	28,2%
Canada	3,5%	13,2%	21,9%	0,9%	6,7%	13,0%	1,9%	8,2%	18,4%
China	0,0%	5,6%	14,2%	1,1%	7,1%	12,6%	0,9%	9,6%	25,6%
Denmark	8,5%	26,5%	39,5%	1,6%	2,4%	2,8%	0,0%	2,3%	8,2%
Egypt	4,0%	15,3%	22,6%	1,2%	3,7%	6,1%	5,1%	13,5%	21,6%
Hungary	1,8%	9,3%	21,1%	3,0%	5,3%	10,2%	5,9%	15,2%	23,1%
Iceland	18,6%	36,9%	46,7%	8,4%	22,3%	30,1%	41,3%	56,4%	64,3%
Israel	9,6%	28,1%	40,3%	2,4%	7,4%	11,1%	0,0%	5,8%	13,7%
Japan	14,6%	24,5%	33,9%	1,5%	11,3%	26,1%	17,4%	30,1%	37,5%
South Korea	2,6%	8,2%	14,3%	2,3%	8,8%	13,8%	15,5%	27,8%	38,6%
Morocco	0,4%	0,9%	2,3%	1,6%	4,7%	7,7%	0,0%	0,5%	4,2%
Norway	16,1%	39,6%	51,9%	2,7%	13,6%	22,2%	15,1%	23,1%	27,9%
Poland	1,6%	5,8%	12,3%	3,9%	9,3%	14,3%	2,9%	8,5%	16,1%
Sweden	7,6%	26,4%	40,4%	2,6%	8,7%	25,4%	2,6%	12,8%	28,8%
Switzerland	10,2%	37,1%	49,7%	3,1%	13,0%	20,9%	13,8%	27,2%	34,2%
Tunisia	6,0%	30,4%	45,5%	1,7%	6,1%	9,1%	5,5%	16,5%	24,2%
Turkey	3,8%	10,3%	14,1%	2,2%	5,3%	8,5%	28,8%	45,8%	55,4%
United Kingdom	9,1%	27,3%	40,1%	1,0%	6,8%	16,4%	7,1%	24,3%	36,7%
United States	3,7%	6,3%	8,3%	1,9%	7,5%	12,7%	7,3%	15,0%	20,2%

Table 2 Percentages of significant area: PPI

	Powe	er - PPI infl	ation	Coherence (PPI-EXCH)		
	SIG01	SIG05	SIG10	SIG01	SIG05	SIG10
Canada	5,7%	11,5%	16,3%	2,6%	13,4%	20,0%
Denmark	7,2%	9,8%	11,7%	0,1%	2,9%	6,3%
Hungary	2,4%	6,9%	11,4%	17,6%	29,0%	36,3%
Japan	16,5%	41,5%	54,9%	34,1%	50,3%	59,1%
South Korea	11,6%	22,4%	31,0%	24,1%	39,1%	49,5%
Norway	0,0%	2,0%	6,0%	1,7%	7,6%	15,0%
Poland	6,6%	21,0%	33,6%	32,4%	46,7%	55,0%
Sweden	1,8%	17,1%	31,0%	20,4%	47,8%	64,1%
Switzerland	20,1%	37,8%	49,7%	8,7%	27,2%	38,3%
Turkey	3,2%	11,2%	18,7%	38,7%	61,4%	72,8%
United Kingdom	14,4%	42,3%	57,8%	11,9%	24,3%	31,5%
United States	7,0%	13,2%	18,8%	2,3%	12,6%	19,7%

The same measure with respect to the PPI (Table 2) shows higher coherency in general, now there are 5 countries that reach the 50 % threshold. Turkey is joined by Sweden, Japan, Poland and South-Korea. (We have no data for Iceland.) This result makes intuitive sense as Sweden and Poland has close trade

relationships with Germany, while Japan and South Korea are substantial industrial exporters, and the PPI and export prices have always been found to move more in accordance with exchange rates than consumer prices.

Dissimilarity measures

The dissimilarity measures (showing the distance between the wavelets) are reported in Table 2. According to the second column the most similar wavelets belong to Sweden, followed by Algeria and Iceland for the CPI. The most dissimilar case is the US. Better indicators are possibly the p-values shown in the third column. These are based on the null hypothesis of independence, and the additional assumption that inflation differentials and log exchange rate changes follow certain ARMA processes. The list of countries where the null is rejected at the 10 % level at least in one of tests is the following: Algeria, Iceland, Japan, Norway, Sweden, Tunisia.

Table 2 Dissimilarity measures

COUNTRY	DISSIMILARITY	P_VAL
Algeria	0,2223	0,054
Canada	0,4876	0,725
China	0,4235	0,58
Denmark	0,4674	0,467
Egypt	0,3322	0,168
Hungary	0,5277	0,516
Iceland	0,2273	0,032
Israel	0,5745	0,81
Japan	0,2820	0,187
South Korea	0,3893	0,344
Morocco	0,5583	0,919
Norway	0,2872	0,076
Poland	0,3746	0,145
Sweden	0,1899	0,006
Switzerland	0,4763	0,682
Tunisia	0,2757	0,084
Turkey	0,4244	0,3
United Kingdom	0,3371	0,293
United States	0,5642	0,861

With respect to the PPI similarity is highest in the cases of Sweden, the UK, and South Korea, but now it is Poland that shows the largest dissimilarity. P-values show that independence is refuted most strongly in the cases of Japan, Sweden, South Korea and the UK.

COUNTRY	DISSIMILARITY	P_VAL
Canada	0,5227	0,772
Denmark	0,4477	0,106
Hungary	0,5783	0,921
Japan	0,2657	0,05
South Korea	0,2268	0,015
Norway	0,5043	0,704
Poland	0,6274	0,92
Sweden	0,2097	0,012
Switzerland	0,4023	0,379
Turkey	0,3615	0,159
United Kingdom	0,2246	0,029
United States	0,5790	0,83

Maximal average coherency: scale³

Regarding the CPI one can see that looking for the scale with maximal average coherency the result is extreme in most cases. When we maximize over frequencies from 1 to 10 years in much more than 50 % of cases the maximum is achieved at either the 1 or 10 years scales. When we restrict maximization to 1 to 6 years, 6 years appears frequently. In the cases where the maximum is achieved at 1 year the overall coherency is rather small. The situation with respect to the PPI is essentially the same. The longer the scale the coherency (if any) seems to be stronger.

Table 3: CPI maximum coherencies

	MAX AVG	MAX AVG	MAX AVG	
COUNTRY	COHERENCE	COHERENCE	COHERENCE	
	(PERIOD 1-10)	(PERIOD 1-6)	(TIME)	
Algeria	8,69	5,90	1994M02	
Canada	9,99	5,90	1991M08	
China	9,99	3,58	2007M12	
Denmark	1,00	1,00	1992M04	
Egypt	4,11	4,11	2017M05	
Hungary	9,99	3,20	2004M07	
Iceland	9,99	5,90	2009M01	
Israel	9,99	1,00	1999M07	
Japan	9,99	5,90	2017M05	
South Korea	9,99	5,90	1998M05	
Morocco	6,23	5,90	2017M05	
Norway	6,96	5,90	2017M05	
Poland	3,48	3,48	1991M08	
Sweden	9,99	1,28	1991M08	
Switzerland	9,99	5,90	2011M04	
Tunisia	9,99	5,90	1997M04	
Turkey	6,23	5,90	1994M06	
United Kingdom	9,99	3,89	1997M12	
United States	1,00	1,00	1991M08	

Table 4 PPI maximum coherencies

³ The Appendix visualizes the following tables.

COUNTRY	MAX AVG COHERENCE (PERIOD 1-10)	MAX AVG COHERENCE (PERIOD 1-6)	MAX AVG COHERENCE (TIME)
Canada	1,00	1,00	2012M12
Denmark	5,90	5,90	2008M11
Hungary	9,99	4,59	2009M02
Japan	9,99	5,90	1998M05
South Korea	9,99	5,90	1997M05
Norway	6,41	5,90	2008M06
Poland	9,99	5,90	2009M03
Sweden	8,46	5,90	1991M08
Switzerland	9,99	5,90	2014M10
Turkey	9,99	5,90	1994M06
United Kingdom	9,99	5,74	1991M11
United States	9,99	5,90	2008M06

Maximal average coherency: time

One can guess that the time index with the maximum average coherency may be around 2008, as it is frequently thought that the connection between exchange rates and prices may strengthen when significant shocks hit the economy. In fact, concerning the CPI nothing similar is observable in our statistics. The most common date is 1991-1992, but this can be due to the fact that these observations are on the edge of our series. However, if we look at the PPI figures, there 2008 seems to be a more prominent time. Of the 12 countries 5 had the highest average coherency either in 2008 or in 2009.

4 Conclusions

In this paper we investigated the relationship between the nominal exchange rate and international relative prices, taking Germany as the country of reference. Our methodology was based on the wavelet and cross-wavelet transforms that posit weaker requirements on the data generating process than traditional time-domain methodologies. Our results are in, some ways, accordance with former findings, but suggest an even gloomier view on the possibility of finding statistically solid relationships between exchange rates and aggregate price indices (CPI or PPI).

In line with the literature we haven't found strong co-movement between prices and exchange rates in the short or medium term, and though coherence seems to be increasing with time scale, it seems far from being stable. For instance, we haven't found any perceivable relationship with respect to Denmark, a country that *a priori* must have strong connection with Germany.

There are weak indications that at least in some countries the price-exchange rate nexus strengthened during the crisis, but even this finding is rather partial and uncertain. In any case detectable cycles at business cycle frequencies do not appear at all.

Though by and large the lead-lag relationship between prices and exchange rates is the expected one, still this is unstable practically for every country, to provide a basis for reliable forecasting.

As we have expected results with respect to the PPI are more promising, but very similar features could be observed as in the case of the CPI. The three countries that seem to be closest to theoretical expectations are Sweden, Japan and South-Korea, where the first should be one of the foremost candidates for this position, while the latter two not. Indeed, this may indicate that the coherence between exchange rates and prices on the macro level may depend more on similarities of export

structure than on trading relations, that the microeconomic intuition would suggest. Also, it is possible that macro price indices are too noisy by their very nature to be amenable for statistical analyses not relying on strong presumptions.

5 Literature

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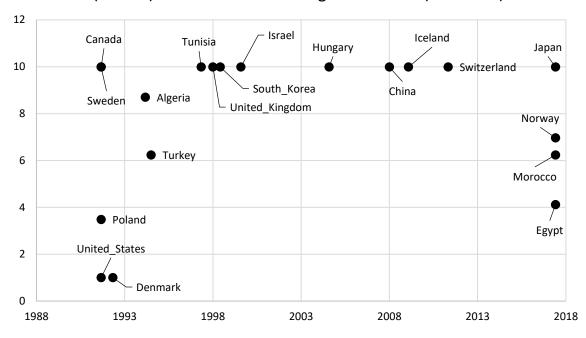
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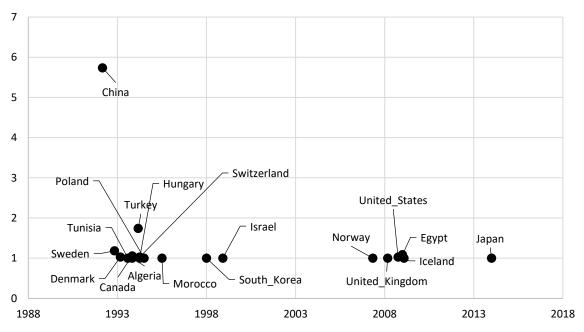
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Appendix: Wavelet Power Figures

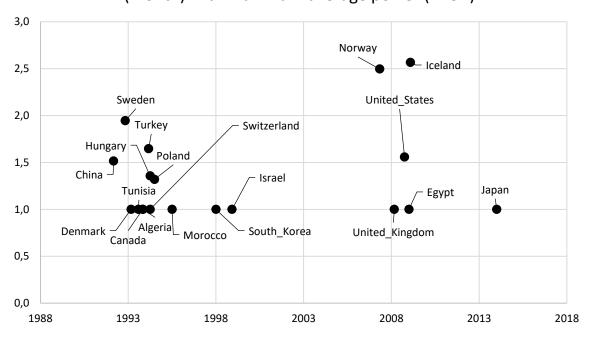
Period (in years) with maximum average coherence vs. time (month) with maximum average coherence (CPI-EXCH)



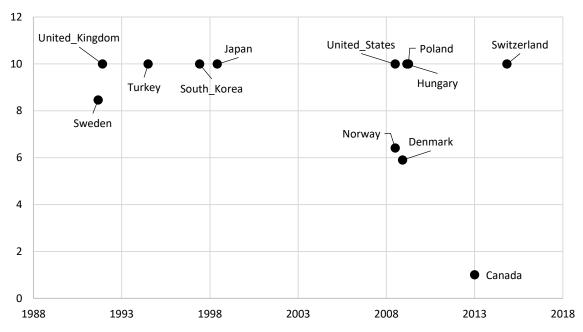
Period (in years) with maximum average power vs. time (month) with maximum average power (CPI)



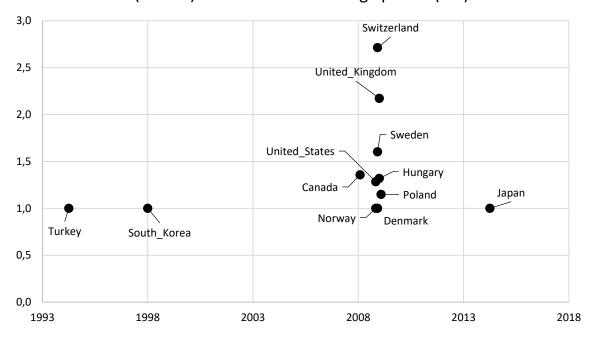
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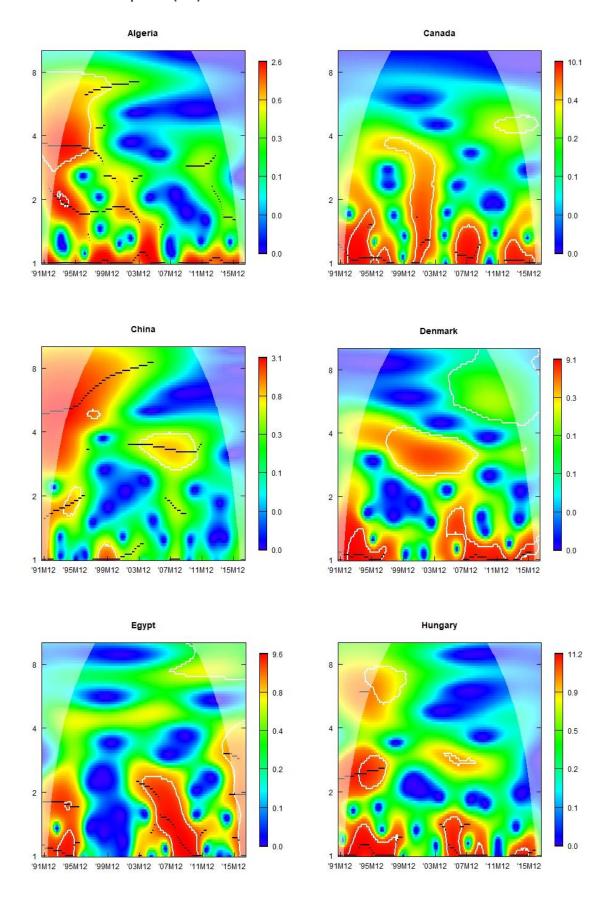
Period (in years) with maximum average coherence vs. time (month) with maximum average coherence (PPI-EXCH)

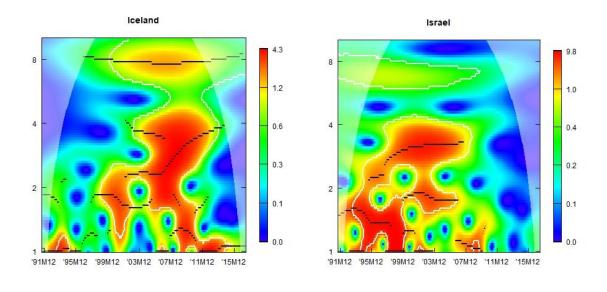


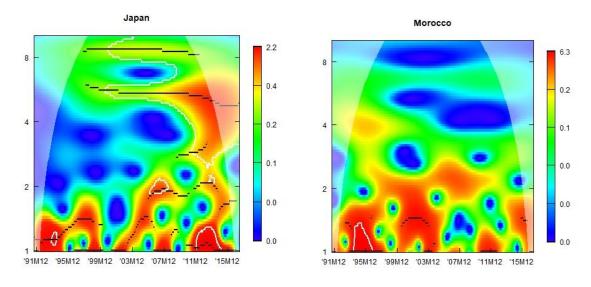
Period (in years) with maximum average power vs. time (month) with maximum average power (PPI)

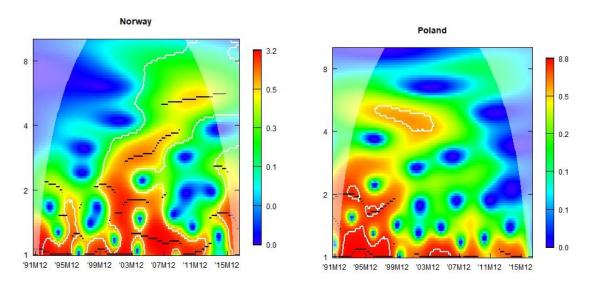


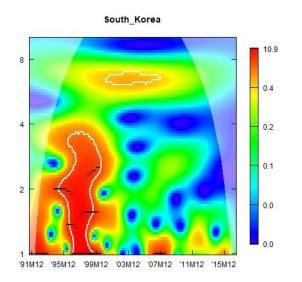
FIGURES -> Wavelet power (CPI)

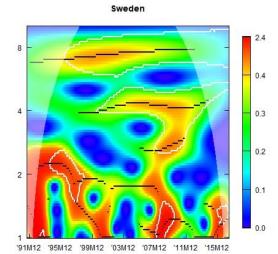


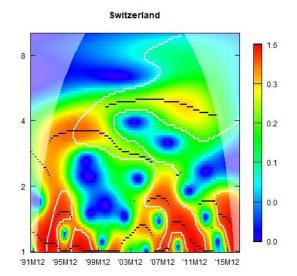


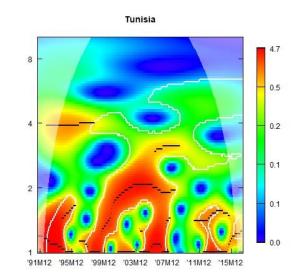


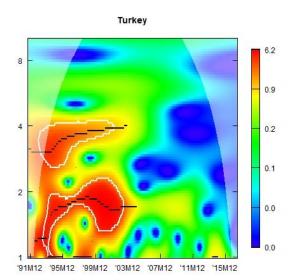


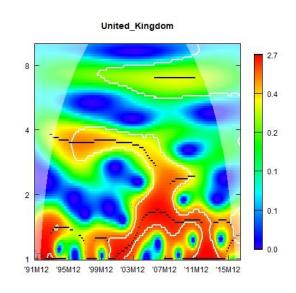




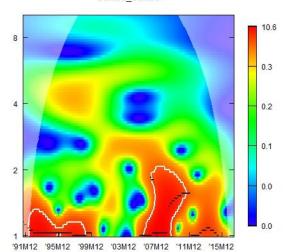




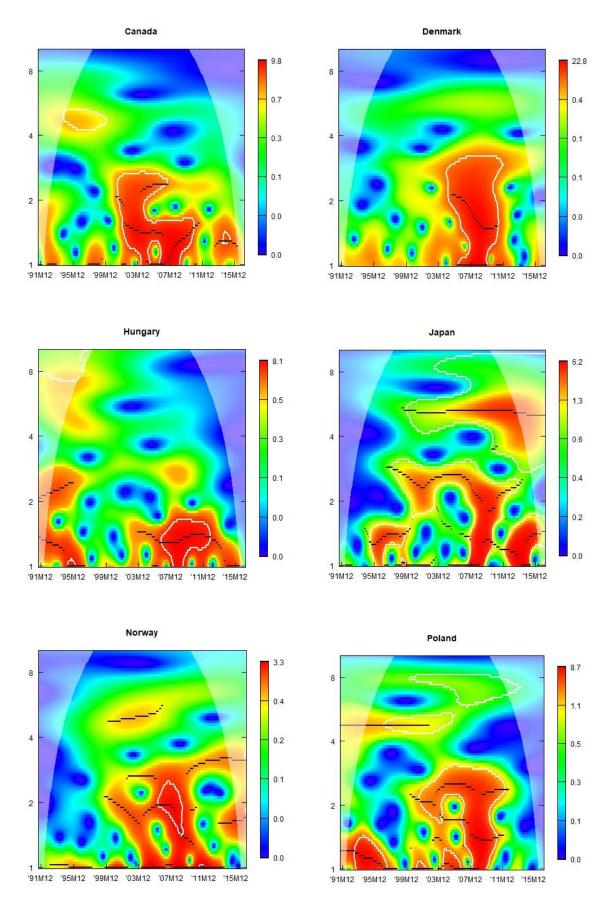


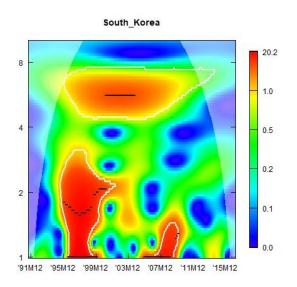


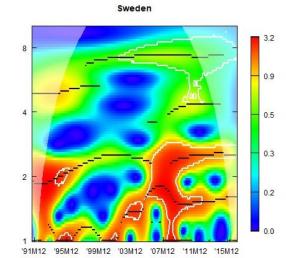
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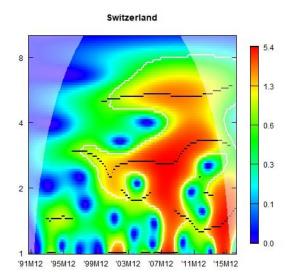


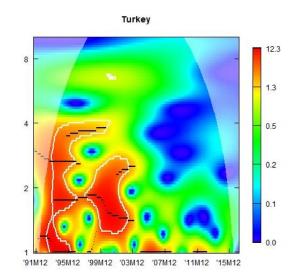
FIGURES -> Wavelet power (PPI)

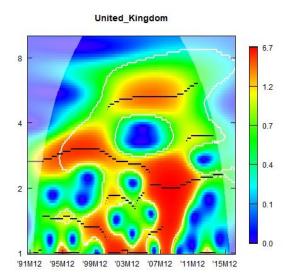


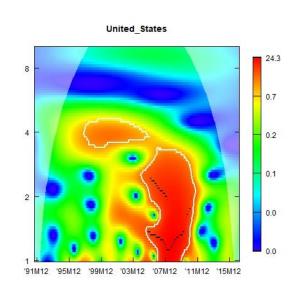




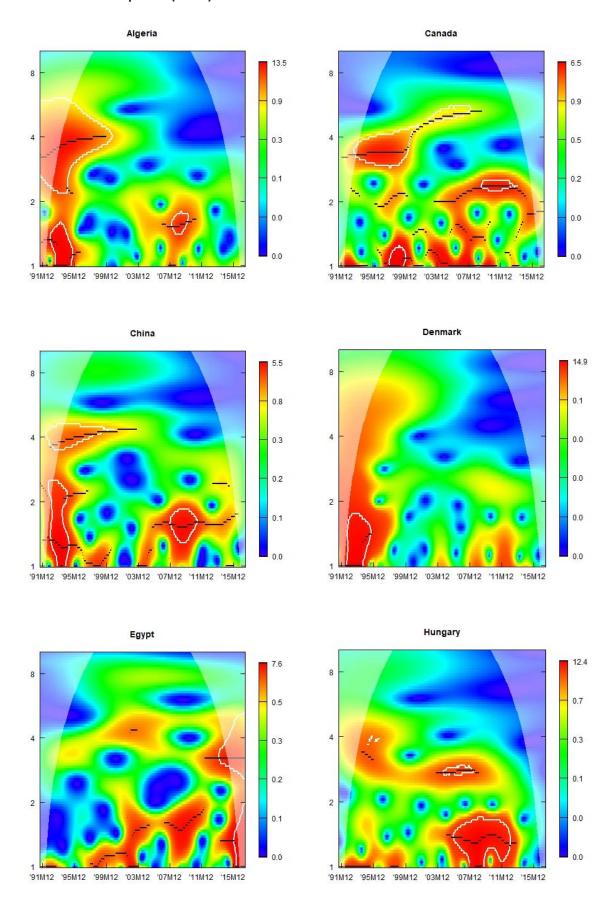


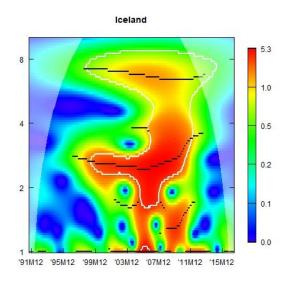


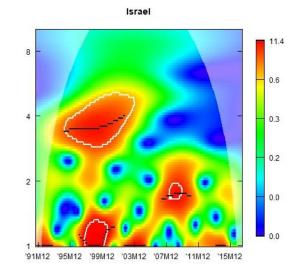


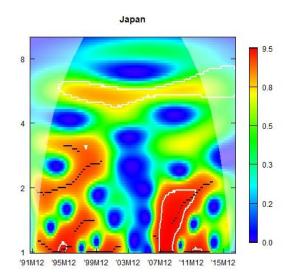


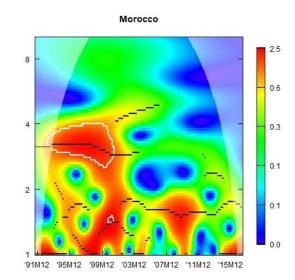
FIGURES -> Wavelet power (EXCH)

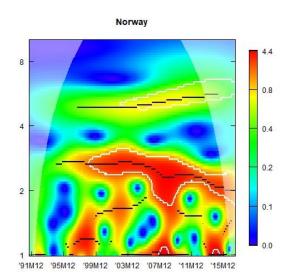


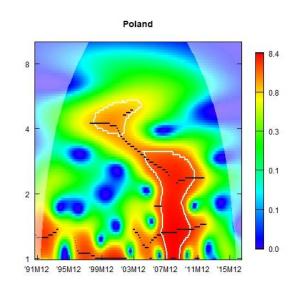


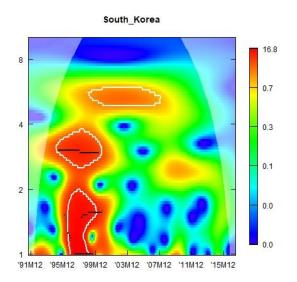


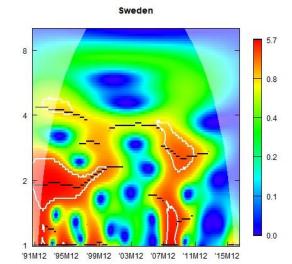


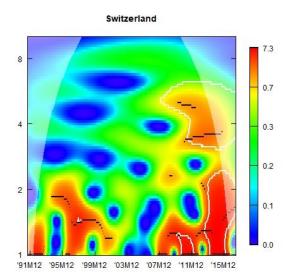


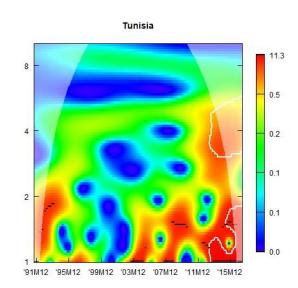


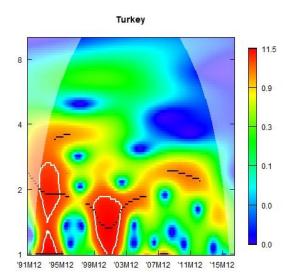


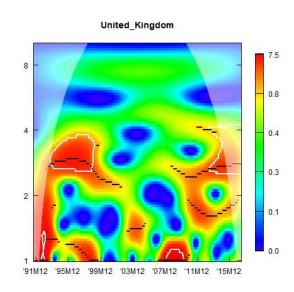












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