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Working Paper

The Intellectual Spoils of War? Defense R&D, Productivity and International Spillovers

IZA Discussion Papers, No. 12769

Provided in Cooperation with:

IZA – Institute of Labor Economics

Suggested Citation: Moretti, Enrico; Steinwender, Claudia; Van Reenen, John (2019) : The Intellectual Spoils of War? Defense R&D, Productivity and International Spillovers, IZA Discussion Papers, No. 12769, Institute of Labor Economics (IZA), Bonn

This Version is available at:

<https://hdl.handle.net/10419/215165>

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Spillovers**

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NOVEMBER 2019

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The Intellectual Spoils of War? Defense R&D, Productivity and International Spillovers

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ISSN: 2365-9793

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ABSTRACT

The Intellectual Spoils of War? Defense R&D, Productivity and International Spillovers*

In the US and many other OECD countries, expenditures for defense-related R&D represent a key policy channel through which governments shape innovation, and dwarf all other public subsidies for innovation. We examine the impact of government funding for R&D - and defense-related R&D in particular - on privately conducted R&D, and its ultimate effect on productivity growth. We estimate models that relate privately funded R&D to lagged government-funded R&D using industry-country level data from OECD countries and firm level data from France. To deal with the potentially endogenous allocation of government R&D funds we use changes in predicted defense R&D as an instrumental variable. In both datasets, we uncover evidence of “crowding in” rather than “crowding out,” as increases in government-funded R&D for an industry or a firm result in significant increases in private sector R&D in that industry or firm. A 10% increase in government-financed R&D generates 4.3% additional privately funded R&D. An analysis of wages and employment suggests that the increase in private R&D expenditure reflects actual increases in R&D employment, not just higher labor costs. Our estimates imply that some of the existing cross-country differences in private R&D investment are due to cross-country differences in defense R&D expenditures. We also find evidence of international spillovers, as increases in government-funded R&D in a particular industry and country raise private R&D in the same industry in other countries. Finally, we find that increases in private R&D induced by increases in defense R&D result in significant productivity gains.

JEL Classification: O30

Keywords: agglomeration, spillovers

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* We thank the Economic and Social Research Council for their financial support through the Centre for Economic Performance. Patrick Warren has provided truly outstanding research assistance. Pierre Azoulay, Josh Lerner, Heidi Williams, and participants in many seminars have given helpful comments. Mirko Draca kindly made his data on US defense procurements available to us, which was invaluable. We are also grateful to Sharon Belenzon and David Thesmar for useful discussions and sharing their data with us.

1 Introduction

We study the impact of government funding for R&D on privately conducted and financed R&D, and its ultimate effect on productivity growth. We focus on an important but relatively understudied component of public policy on R&D: defense-related R&D. Defense R&D represents a key channel through which governments all over the world shape innovation. In the US, annual government defense-related R&D expenditures amount to about \$78.1 billion in 2016, 57.2% of all government-funded R&D (Congressional Research Service, 2018). While defense-related R&D is motivated by goals that are not mainly economic, it is often the most important *de facto* industrial policy used by the federal government to affect the speed and direction of innovation in the economy. The amount of public money flowing into defense R&D dwarfs the amount spent on other prominent innovation policy tools in the US. For example, the total budget of the National Science Foundation or the overall value of the federal R&D tax credit in a typical year are less than one tenth of federal outlays for defense-related R&D (NSF 2006). Defense R&D is the single most important component of government-funded R&D in the UK and France as well, and a major component of government-sponsored R&D in many other developed economies.

In this paper we use two complementary panel datasets—a country-industry-year-level dataset for OECD countries and a firm-year-level dataset for France—to address two related questions. First, we estimate the effect of government-funded R&D on private R&D—namely, R&D conducted and financed by private businesses. We are interested in whether government-funded R&D in a given country and industry (or to a given firm) displaces or fosters private R&D in the same country and industry (or firm). We use arguably exogenous variation in defense-related R&D to isolate the causal effect of government-funded R&D. Having found evidence of a positive effect, we next estimate how investment in R&D affects productivity. For both types of analysis, we assess whether the benefits of public R&D investment are limited to a single country or spill over across multiple countries.

The effect of defense R&D expenditures on private sector innovation and economic growth has been a hotly debated topic for many years (e.g., see surveys by Mowery, 2010, and Lichtenberg, 1995). Proponents of the benefits of defense R&D point to the commercial success of major innovations such as jet engines, computers, radar, nuclear power, semiconductors, GPS, and the Internet as evidence that

military R&D has been a crucial source of technological development with civilian applications.¹ Some even argue that the Pentagon's role as the world's most generous investor in technological innovation during the Cold War—ultimately resulting in superior technologies for American companies and enduring gains in their competitiveness (Braddon, 1999)—was an important reason that US manufacturing became so dominant after World War II. More recently, defense R&D has been viewed as an important contributor to national economic growth through private sector spinoffs and agglomeration economies.² Proponents of this view often point to Israel as an example of how defense spending has spawned a multitude of commercially successful high tech startups (e.g. Senor and Singer, 2009).

On the other hand, critics argue that the social returns to defense R&D are likely to be low because the secrecy that surrounds defense R&D inherently limits the scope of spillovers to civilian firms. Even more fundamentally, critics argue that defense-related R&D might *displace* private R&D and therefore could have little (or even a negative) impact on the total amount of innovation in a country. Overall, there is much anecdotal evidence of some of the positive and negative effects that defense R&D might have on growth, but little systematic econometric evidence.³

We begin our empirical analysis using a unique dataset that we constructed by linking detailed information on defense-related and non-defense related government-funded R&D to information on private R&D, output, employment, and salaries in 26 industries in all OECD countries over 23 years. We estimate models that relate privately funded R&D in a given country, industry, and year to

¹ For example, see Lichtenberg (1984, 1988), Ruttan (2006), Mazzucato (2013) and the recent discussion of the DARPA model by Azoulay et al. (2019b). Draca (2013) estimates the impact of US defense spending on firm-level innovation and finds that increases in procurement contracts are associated with increases in patenting and R&D.

² An additional benefit of military R&D is the creation of highly specialized human capital valued by the private sector. Silicon Valley companies increasingly scout Pentagon and NSA personnel for potential hires (Sengupta, 2013)

³ For a recent survey of the literature on the evaluation of innovation policies, see Bloom, Van Reenen, and Williams (2019). The literature focuses on two types of R&D policies. First, there are fiscal policies towards R&D such as Hall (1993), Bloom, Griffith, and Van Reenen (2002), Moretti and Wilson (2014), Dechezlepretre et al. (2016) and Rao (2016). This generally finds positive effects. Second, there is a body of empirical research on the effect of public R&D on private R&D (e.g. David, Hall, and Toole, 2000; Lach, 2002; Goolsbee, 1998; Wallsten, 2000). The results here are mixed – see for example the meta-study of Dimos and Pugh (2016). There is a small number of recent papers that have clearer causal identification strategies, but these papers tend to differ from ours in that they focus on single countries, different policy instruments, and different outcomes than those we focus on. For example, Jacob and Lefgren (2011) and Azoulay et al. (2019a) study the effect of NIH grants on publications and patenting; Bronzini and Iachini (2014) focus on the effect of R&D subsidies on capital investment by Italian firms; Howell (2017) looks at Department of Energy SBIR grants on venture capital funding and patents; and Slavtchev and Wiederhold (2016) study the effect of government procurement on innovation. Guellec and van Pottelsberghe de la Potterie (2001) and Pless (2019) are rare exceptions that look at multiple types of R&D policies.

government-funded R&D in the previous year, conditioning on a full set of country-industry and industry-year fixed effects.

We complement this industry-level analysis with a firm-level analysis based on a longitudinal sample of firms that engage in R&D followed collected by the French Ministry of Research from 1980 to 2015. This is the only available dataset we know of that disaggregates public defense R&D subsidies by firms across the whole economy. One advantage of using firm-level data is that we observe which firms within an industry actually receive public R&D funds and which do not. The longitudinal nature of the data allows us to control for firm fixed effects, absorbing all time invariant unobserved differences across firms that may be systematically correlated with the propensity to invest in R&D. We compare the same firm to itself in different moments in time and identification stems from the exact timing of the public R&D award.

In both the OECD data and the French data, we use *predicted* defense R&D as an instrumental variable to isolate exogenous variation in public R&D. This instrument combines nationwide changes to defense R&D with fixed allocations across industries. Predicted defense R&D provides arguably exogenous variation because annual aggregate changes in defense spending reflect political and military priorities that are largely independent of productivity shocks in different domestic industries. Wars, changes of government, and terrorist attacks have had major influences on defense spending. In the US, for example, military R&D spending ramped up under Reagan; fell back after the end of the Cold War, and rose again after 9/11 (we use the 9/11 shock as an econometric case study of our method). Importantly for our identification strategy, the impact that nationwide exogenous changes in military spending have on defense related R&D varies enormously across industries, because some industries (e.g. aerospace) always rely more heavily on defense-funding than others (e.g. textiles).⁴

The sign of the effect of government-funded R&D on privately-funded R&D could be positive or negative, depending on whether government-funded R&D crowds out or crowds in privately-funded R&D. Crowding out may occur if the supply of inputs to the R&D process (specialized engineers, for example) is inelastic within an industry and country (Goolsbee, 1998). In this case, the only effect of an increase in government-funded R&D is to displace private R&D with no net gains for total R&D.

⁴ The idea of using military spending as an exogenous component of government spending has been used in other contexts. In the analysis of fiscal multipliers Ramey (2011) and Barro and Redlick (2011) have argued for the importance of using defense spending to mitigate endogeneity concerns. See also Perotti (2014).

Crowding in may occur if (i) R&D activity involves large fixed costs and, by covering some of the fixed costs, government-funded R&D makes some marginal private sector projects profitable;⁵ (ii) government-funded R&D in an industry generates technological spillovers that benefit other private firms in the same industry; and/or (iii) firms face credit constraints.

Empirically, we find strong evidence of crowding-in in both the OECD and French datasets. Increases in government-funded R&D generated by variation in predicted defense R&D translate into significant increases in privately-funded R&D expenditures, with our preferred estimates of the elasticity equal to 0.43. Our estimate implies that defense-related R&D is responsible for an important part of private R&D investment in some industries. For example, in the US “aerospace products and parts” industry, defense-related R&D amounted to \$3,026 million in 2002 (nominal). Our estimates suggest that this public investment results in \$1,632 million of additional private investment in R&D. Our estimates also indicate that cross-country differences in defense R&D might play an important role in determining cross-country differences in overall private sector R&D investment. For example, we estimate that if France increased its defense R&D to the level of the US as a fraction of its GDP (admittedly a large increase) private R&D in France would increase by 8.7%.

The increases in private R&D expenditures appear to reflect actual increases in R&D activity, not just higher wages and input prices caused by increased demand. We uncover significant positive effects on employment of R&D personnel in both datasets, with limited wage increases. This is consistent with a fairly elastic local supply of specialized R&D workers within an industry across countries, or across industries.

We find evidence of spillovers between countries.⁶ In the OECD data, increases in government funded R&D in one country appear to increase private R&D spending in the same industry in other countries. For example, an increase in government-funded R&D in the US chemical industry induced by an increase in US defense spending in the chemical industry raises the industry’s private R&D in the US, but it also raises private R&D in the German chemical industry. This type of cross-border spillover is consistent with the presence of industry-wide technological or human capital externalities.

⁵ Examples of fixed costs include labs that can be used both for government-financed R&D and for private R&D or human capital investment in the form of learning by scientists on topics that have both military and civilian applications.

⁶ International spillovers of R&D are studied by Hall, Mairesse, and Mohnen (2010); Coe and Helpman (1995); Pottelsberghe and Lichtenberg (2001); Keller (2004); and Bilir and Morales (2015).

In the final part of the paper, we turn to the effect of investment in R&D on productivity. We estimate models where measures of productivity growth – growth in TFP or output per worker - are regressed on lagged private R&D intensity, using predicted defense R&D as a share of value added or sales as an instrumental variable. Industry-level models based on OECD data indicate a positive effect of private R&D on TFP. Quantitatively, the return to R&D appears to be economically meaningful and confirms the key role played by innovation in driving economic growth.⁷ OLS estimates from firm-level models based on French data are qualitatively consistent with these findings, but IV estimates in this set of models are very noisy.

An increase in the defense R&D to value added ratio of one percentage point is estimated to cause a 5% increase in the yearly growth rate of TFP (e.g., from 2 percent per annum to 2.1 percent). We view this as an important, but not overwhelming effect. It suggests that a fraction, although certainly not all, of US economic growth is accounted for by investment in defense R&D. For example, defense R&D in the US increased by 52% between 2001 and 2004 following the 9/11 attack. Compared to historical standard, this was a very large increase. We estimate that, holding taxes constant, this translated into a 0.006 percentage point increase of the annual TFP growth rate in the US in the affected years—about a 1.8% increase. Since in reality, this defense R&D spending had to be financed by increased taxes or cuts in other government expenditures, the ultimate impact is almost certainly smaller.

Overall, our estimates suggest that cross-country differences in defense R&D play a role in explaining cross-country differences in private R&D investment, speed of innovation, and ultimately in productivity of private sector firms.

We caution that our estimates do not necessarily imply that it is desirable for all countries to raise defense R&D across the board. Our finding that government-funded R&D results in increased private R&D does not necessarily imply that defense R&D is the most efficient way for a government to stimulate private sector innovation and productivity. There are other possible innovation policies available to governments (Bloom, Van Reenen and Williams, 2019) and our analysis does not compare defense R&D to other types of public R&D spending.

⁷ Consistent with the existence of international technology spillovers, we also uncover a positive effect of investment in R&D in an industry and country on productivity of firms in the same industry but in different countries.

The structure of the paper is as follows. Section 2 presents a simple framework and the empirical models. Section 3 describes the data. Sections 4 and 5 present the empirical results. Section 6 concludes.

2 Conceptual Framework, Econometric Models and Identification

We start with a simple framework that is useful in deriving the empirical models we take to the data, and in clarifying how to identify and interpret our empirical estimates. Here we focus on the effect of government-funded R&D on private R&D activity. Specifically, we are interested in the direct effect for an industry-country pair or a firm of receiving government-funded R&D on the recipient's own private R&D investment. In addition, we are interested in international spillover effects that might arise if changes in government-funded R&D in a particular industry and country indirectly affect private R&D activity in the same industry in other countries. In Section 5, we will discuss the framework to estimate the direct and indirect effects of R&D investment on productivity.

2.1 Conceptual Framework

We assume that output of firm f in industry i in country k at time t is a function of capital, K , labor, L , and intermediate inputs, M :

$$Y_{fikt} = A_{fikt}F(K_{fikt}, L_{fikt}, M_{fikt}) \quad (1)$$

where A is Hicks-neutral TFP. Following the R&D literature (e.g. Griliches, 1979) we assume that A is determined by the lagged R&D knowledge stock G :

$$\ln A_{fikt} = \eta \ln G_{fik(t-1)} + \gamma X_{fikt} + u_{fikt} \quad (2)$$

where $\eta = \frac{\partial Y}{\partial G} \frac{G}{Y}$ is the elasticity of output with respect to the business R&D stock, X are other factors influencing TFP, and u_{fikt} is a stochastic error term. The R&D stock G is an increasing function of privately funded R&D expenditures (R) and government-funded R&D expenditures (S).

We assume that S is set by the government, while R is chosen by the firm based on the technology embodied in equations (1) and (2), as well as on the cost of R&D. We assume that the (static) demand for private R&D can be written as:

$$\ln R = \sigma \ln U(S) + \beta \ln Y + \nu \quad (3)$$

where U is the Hall-Jorgenson tax-adjusted user cost of R&D capital, which is allowed to depend on public subsidies, S (see Criscuolo et al, 2019, for a discussion of how subsidies affect the user cost of capital). The user cost will also depend on current and expected interest rates, depreciation and the tax system as a whole (including R&D tax credits). This equation can be rationalized as the steady state demand for R&D from the first order conditions from specializing equation (1) to a CES production function (e.g. Bloom, Griffith, and Van Reenen, 2002). Under this interpretation, σ is the elasticity of substitution and β is the returns to scale parameter ($\beta = 1$ indicates constant returns) and ν is a function of technological parameters in the production function indicating factor biases.⁸

In our empirical analysis, we use two alternative datasets to estimate two variants of equation (3). In our analysis of OECD data, the level of observation is an industry-country-year and we assume that we can take a first order approximation of equation (3) as:

$$\ln R_{ikt} = \alpha^{OECD} \ln S_{ik(t-1)} + \beta^{OECD} \ln Y_{ikt} + \lambda X_{kt} + d_{ik} + d_{it} + v_{ikt} \quad (4)$$

where the determinants of R&D other than S and Y are assumed to be a vector of country by year observables X_{kt} , a set of industry by country fixed effects (d_{ik}), industry by year dummies (d_{it} , e.g. industry specific product demand or technological shocks), and an idiosyncratic error (v_{ikt}).⁹ In our baseline models, X_{kt} includes current and past GDP levels, thus controlling for country-specific business cycles as these demand side effects are likely to affect innovation (e.g. Shleifer, 1986).

In our analysis of the French data, the level of observation is firm-years, and we assume that we can take a first order approximation of equation (3) as:

$$\ln R_{fit} = \alpha^{FRA} \ln S_{fi(t-1)} + \beta^{FRA} \ln Y_{fit} + d_f + d_t + v_{fit} \quad (5)$$

⁸ If the production function was more general than CES, ν would also include other factor prices such as the wage rate.

⁹ We include dummy variables to control for the country-industry fixed effects, which formally requires strict exogeneity of the right hand side variables. Since our panel is long (up to 26 years) we do not think there is likely to be much bias from this issue, but to check we also estimated the equation in first differences, which requires weaker exogeneity assumptions and obtained similar results.

where we include a set of firm fixed effects (d_f) to absorb all sources of time-invariant heterogeneity across firms. Since in this specification we only have one country, we do not include X_{kt} and d_{ik} - these are absorbed by the time dummies and firm fixed effects respectively. We also consider industry-level versions of equation (5) where we can disaggregate industries at a finer detail than in the OECD data.

Equations (4) and (5) represent our baseline models in the empirical analysis of the effects of government R&D on private R&D. The focus of our analysis is on estimating the coefficients α^{OECD} and α^{FRA} that relate changes in government-funded R&D in a given year to changes in private R&D in the following year. (Recall that the dependent variable is only the privately funded part of business R&D, so there is no mechanical association between R and S .)

The sign of α is unknown a priori. If increases in government-funded R&D crowd *out* private R&D, the α terms should be negative. In the case of complete crowding out, the only effect of the policy is to displace private R&D, with no net gain in total R&D. This would be the case if, for example, the supply of inputs in the R&D process in any given industry was perfectly inelastic in the short run. A key input in this respect is likely to be specialized scientists and engineers and the elasticity of their supply to a country-industry depends on their mobility across industries and countries. With inelastic supply to a country-industry, increases in public funds for R&D come at the expense of declines in private R&D.

If, on the other hand, increases in government funded R&D crowd *in* private R&D, the α terms should be positive. In this case, more public R&D stimulates even more private R&D. There are at least three possible reasons for why this might be the case.

First, in the presence of large fixed costs, public R&D may make marginal private projects feasible. In most industries, R&D activity is characterized by large fixed costs in the form of labs, research, human capital accumulation, set up costs, etc. It is realistic to think that some of these fixed costs can be used for multiple projects. For example, lab infrastructure set up for a specific project can, in some cases, be used for other projects as well. Similarly, a scientist's human capital acquired while working on a specific project—the intellectual understanding of a specific literature, for example, or her mastery of a scientific technique—can be helpful in other projects. By paying for some of the fixed costs, government-funded R&D may make profitable for private firms' projects that otherwise would not have

been profitable. Similarly, if government-funded R&D results in process innovation, it is conceivable that this innovation can indirectly benefit private R&D.

Second, if firms are credit constrained, the public provision of R&D might relax these financial constraints. Although capital markets are generally well developed for the OECD countries we study, the special nature of R&D investments highlighted by Arrow (1962), such as riskiness and asymmetric information, may make it especially vulnerable to financial frictions.¹⁰

Third, government-financed R&D investment by one firm may make other firms in the same industry more productive because of technology or human capital spillovers (e.g. Moretti, 2004 and 2019). In this case, an increase in government-financed R&D directly raises R&D in the firm that receives the government contract, and may indirectly raise R&D in other firms in the same industry or same locality. Spillovers could also be negative in the case of strategic substitutability, as rival firms could free ride off the R&D of the supported firms (e.g. Bloom, Schankerman, and Van Reenen, 2013).

An implication is that in the presence of R&D spillovers within an industry the estimated coefficient from industry-level data in equation (4) does not need to be identical to the coefficient from firm-level data in equation (5). Broadly, we expect industry coefficients should be larger if crowd-in induces rival firms to do more R&D (due to strategic complementarity) or smaller if rivals do less R&D (due to strategic substitutability).

In terms of inference: to account for the possible correlation of residuals in each year across industries in a given country and across countries in a given industry, standard errors for OECD data throughout the paper are multi-way clustered by country by industry pair and country by year pair (Miller, Cameron and Galbech, 2009). In the regressions based on the French data, we cluster at the 2-digit industry for industry-level regressions; and 3-digit industry for firm-level regressions.

2.2 Identification and Threats to Validity

Equations (4) and (5) allow us to control for a wide variety of shocks that affect private R&D and may also be correlated with government-financed R&D. In equation (4), the inclusion of industry-year dummies accounts for the fact that different industries have different propensities to invest in R&D, and these differences can vary over time as a function of technology shocks and product demand shocks. The inclusion of industry-country fixed effects accounts for differences in the propensity to invest in

¹⁰ See Garicano and Steinwender (2016) for some empirical evidence on financial frictions for R&D.

R&D across countries and the fact that these international differences may be more pronounced in some industries than others.

In equation (5), the vector of controls also include firm fixed effects. This is an important advantage of using firm-level data. It allows us to absorb all time-invariant unobserved differences across firms that may be systematically correlated with propensity to invest in R&D. In equation (5), we compare the same firm to itself at different moments in time. Identification stems from the timing of the public R&D award.

Even after conditioning on this rich set of controls, our models yield inconsistent estimates if the timing of public R&D is correlated with unobserved time-varying determinants of private R&D. We cannot rule out this possibility, as government policies are unlikely to be random. It is possible that governments allocate R&D funding to specific industries or specific firms based on criteria that are correlated with determinants of private R&D investment. This may happen, for example, if governments tend to use public funds to help firms in sectors that are struggling and are experiencing declines in private R&D. In this case, changes in public R&D would be negatively correlated with unobserved determinants of private R&D, introducing a negative bias in our estimates of the coefficient α in equations (4) and (5). The opposite bias arises if governments tend to use public funds to help firms in sectors that are thriving, and are experiencing increases in R&D over and above those experienced by the same sector in other countries.

In either case, $S_{ik(t-1)}$ might be correlated with v_{ikt} in equation (4) and $S_{fik(t-1)}$ might be correlated with v_{fikt} in equation (5). If governments help winners, the correlation between $S_{ik(t-1)}$ and v_{ikt} (and $S_{fik(t-1)}$ and v_{fikt}) is positive. If governments disproportionately help “losers” (compensatory policies), the correlation between $S_{ik(t-1)}$ and v_{ikt} (and $S_{fik(t-1)}$ and v_{fikt}) is negative. Note that in equation (4) what matters are *industry-country specific time-varying* shocks. Equation (4) is robust to industry specific time-varying shocks shared by all countries. For example, if the telecommunication industry is struggling in *all* countries, and governments decide to endogenously increase publicly funded R&D for the industry, equation (4) would yield consistent estimates.

To deal with these issues, we use *predicted* public defense R&D subsidies as instrumental variable for publicly funded R&D. Mowery (2010) explains that defense R&D is the most important example of “mission-oriented R&D,” i.e., R&D that is spent not in order to pursue economic goals, but any other, unrelated government objectives. In particular, defense R&D is usually motivated by

geopolitical, not economic, considerations. Our identifying assumption is that time variation in defense-related R&D is largely driven by events exogenous to country-specific R&D shocks, such as wars, terrorism, geopolitical shocks like the end of the Cold War, and the ideological preferences of the political leaders in power. Defense R&D is by far the largest component of government R&D in many countries, e.g., United States, United Kingdom, and France. This ensures that our instrument has a strong first stage. Defense R&D also causes the biggest variations in public R&D over time, and there is a large variation across countries, ranging from pacifist country like Japan or neutral countries like Austria, to defense-heavy countries like the United States and South Korea.

Predicted defense R&D subsidies (DR_{ikt}^{IV}) is defined as

$$DR_{ikt}^{IV} = share_{i(t-1)}^l \cdot \overline{def}_{kt}$$

in which we distribute aggregate changes in defense R&D spending at the country level \overline{def}_{kt} across industries according to some industry shares, $share_{i(t-1)}^l$. Using predicted instead of actual defense R&D weakens the power of the instrument in the first stage, but it strengthens its validity. In practice, our first stage has good power and is robust to various changes in the assumptions we use to construct the instrument.

In practice, the exact implementation of the IV differs slightly for the OECD and the French dataset due to differences in level of aggregation and variable definitions. The exact details on how we construct the instrument are in Appendix A. For the OECD analysis, \overline{def}_{kt} is country k 's total defense R&D spending in year t ¹¹ and $share_{i(t-1)}^l = share_{i(t-1)}^{US}$ is government defense R&D in industry i as a proportion of all the government defense R&D in the *United States*. Using $share_{i(t-1)}^{US}$ rather than the own country share reduces the risk that the industry distribution of defense R&D subsidies responds to expected country-specific shocks. Using the US data also has the practical advantage that defense R&D subsidy data at the country-industry level by year is not available for most countries.¹²

¹¹ Notice that this includes not only defense related R&D subsidies spent by businesses, but also spend by sectors including e.g., universities (called “government budget appropriations or outlays on R&D” or GBAORD by the OECD). We do not have aggregate numbers for the business sector only component for all OECD countries (only for France). In Table 3 we conduct robustness tests where we control for non-business R&D to make sure our affects are not being driven by defense R&D carried out in government labs or universities.

¹² The US industry shares are given for each fiscal year, so it is partially lagged compared to the other variables. Apart from the US, the defense breakdown by industry is only available in the UK in 1993-2009 so we use the UK specific information for the UK part of the analysis. It is available for the years 1987-2003 for the US (we hold the 2003 share constant for the years after 2003). Our results are robust to dropping either the US or UK or both. We use the France-specific data in the firm-level analysis below.

Notice that it is possible that while the *overall* level of defense spending in a country is orthogonal to the residual v_{ikt} , the industry composition of defense spending may still be correlated with v_{ikt} . This would be the case if, for example, French defense spending declined after the end of the Cold War for exogenous reasons, but the decline was smaller in, say, aerospace, for endogenous reasons. Because we are using US industry share for all countries (other than the UK), this is a problem only to the extent that endogenous adjustments to the industry share reflect unobserved industry-specific time-varying shocks that are shared by the US and the relevant country. Empirically, models that exclude the UK or US or both yield similar estimates. We also obtain similar estimates when we fix industry weights equal to the US industry share at the beginning of sample period—therefore abstracting from any potentially endogenous change in industry share.

In the French analysis, the IV for firm-level models is defined as $DR_{ft}^{IV} = share_{i4} \cdot \widetilde{def}_{i3,t}$ where $share_{i4}$ is the annual share of defense R&D subsidies allocated to firm f 's main four-digit SIC industry (averaged across all years); and $\widetilde{def}_{i3,t}$ is the level of defense subsidies defined at the three-digit industry level *excluding subsidies going to firm f itself in a particular year*, to avoid a mechanical correlation between the IV and instrumented variable. Analogously, in our three-digit industry level analysis for France, we use the three-digit industry defense R&D share and the two-digit industry defense R&D subsidy excluding the subsidy to firm f 's three digit industry: $share_{i3} \cdot \widetilde{def}_{i2,t}$. We tried using the US shares for the French analysis, but the level of coarseness of the aggregation of industry meant that the first stages were weak.

An issue to consider is that, although shocks to military R&D are unlikely to be related to technology shocks, they could in principle signal shocks to future product demand. Under this view, an event such as 9/11 generated a direct increase in military R&D, but also increased current and future demand for military products. In turn, this second channel could stimulate additional private R&D through a demand or market size effect, thus invalidating our instrument. We note that it is not *ex ante* obvious that this is a major issue in our context because historically large increases in government defense procurement are typically targeted toward existing, rather than new technologies, while most R&D is likely to be directed at new technologies.¹³

¹³ This is why many historians like Milward (1977) have argued that wars tend to retard technological change by engendering a more conservative attitude to military procurement.

Empirically, we present four tests intended to probe the sensitivity of our findings to these demand effects. (i) As a first pass, we estimate models that condition on *future* industry output. These models are not our preferred specification, because future output needs to be thought of as an endogenous variable. (ii) We then estimate models that condition on non-R&D military spending and expectations thereof. (iii) We also estimate our baseline model solely on the US, where we have industry-specific R&D and non-R&D public military spending. (iv) Finally, we perform placebo tests based on components of defense spending that are unrelated to R&D subsidies paid to businesses. The idea is that our instrumental variable estimates should not be driven by changes in defense procurement that stimulate demand rather than R&D. Overall, our findings indicate that changes in expected demand are not major sources of bias in our models.

Another issue to consider in interpreting our IV estimates has to do with the possibility that government-funded R&D by country i is set endogenously in response to government-funded R&D by country i 's competitors. For example, an increase in government-funded R&D in, say, the German chemical sector may induce France to increase its own government-funded R&D in the chemical sector. This does not invalidate our estimates—even in an ideal randomized setting where public subsidies are randomly assigned this issue would arise—but it affects their interpretation. In this case, the α parameters should be interpreted as the effect of S on R , after allowing for the endogenous reaction of other countries. This is arguably the parameter of interest for policy. It informs policymakers of what they can expect from a policy change if other countries react.

2.3 Employment and Wages

We also examine the effect of increases in public R&D investment on employment and wages. This is important because an increase in private R&D *expenditures* does not necessarily equal an increase in R&D *activity*. We distinguish between the effect on labor market outcomes of R&D workers and labor market outcomes of non-R&D workers. If the supply of R&D workers is completely inelastic in the short run, increased R&D spending could simply result in higher wages, with little or no effect on employment and innovation (Goolsbee, 1998). On the other hand, if R&D workers can move across industries or across countries so that supply to a specific country and industry is fairly elastic, we might find significant increases in R&D personnel and limited increases in their wages.

The effects on demand for non-R&D personnel in the industry depend on whether R&D generates technologies that substitute for or complement such labor. On the one hand, more R&D in an industry may result in product innovation, higher sales, and therefore more labor demand. On the other hand, process innovation can easily reduce employment by making it easier to produce the same output with fewer labor inputs.

To empirically assess these questions, we estimate models similar to the one in equations (4) and (5) where the dependent variable is the employment of R&D workers, employment of non-R&D workers, and average wages.

2.4 International Spillovers

It is in principle possible that increases in government-funded R&D in an industry in a given country affect private R&D investment by firms in the same industry located abroad. For example, an increase in government-funded R&D in the German chemical industry may reduce private R&D in the French chemical industry. This may be due to strategic reasons, as French firms decide it is not worth competing to catch up with their German rivals (e.g. international R&D is a strategic substitute) or the cost of internationally used industry-specific R&D inputs (e.g. chemical engineers) may be driven up. The opposite spillover effect could also arise if R&D is a strategic complement between countries, so that increased public R&D in Germany results in French firms investing more to keep up in the race; or if there are significant cross-country technological or human capital externalities.

To empirically assess international spillovers, we use our OECD data to estimate models of the form:

$$\ln R_{ikt} = \alpha^{OECD} \ln S_{ik(t-1)} + \gamma^{OECD} \ln SP_{ik(t-1)} + \beta^{OECD} \ln Y_{ikt} + \lambda X_{kt} + d_{it} + d_{ik} + v_{ikt} \quad (6)$$

where $SP_{ik(t-1)}$ is a weighted average of government-funded R&D in other countries in the same industry and year with weights measuring the between country i and each other country: $SP_{ik(t-1)} = \sum_j d_{ij} S_{jk(t-1)}$ where d_{ij} is the economic or geographic “distance” between country i and country j (normalized to sum to one for each country i) and $S_{jk(t-1)}$ is, as before, government-funded R&D in industry i in country j .

3 Data and Basic Facts

We use two separate data sets: section 3.1 details the data used in our OECD industry-country level analysis; section 3.2 describes the data used in our French firm-level analysis.

3.1 OECD Industry-Country Data

Data Sources. We combine data for OECD countries from the SStructural ANalysis (STAN) dataset and the Main Science and Technology Indicators (MSTI) dataset. Our data include 26 countries, 26 industries, and 23 years, from 1987 to 2009. The Data Appendix describes in detail how we cleaned and merged the data and provides the exact definition of each variable with the corresponding source.

The definitions of R&D are based on the internationally recognized “Frascati Manual” used by the OECD and national statistical agencies. Our main R&D variable measures industry-level R&D *conducted by businesses* (known as “Business Enterprise R&D” or “BERD”).

We will generally refer to BERD as simply “R&D” for brevity. While all BERD is conducted by firms, some of its funding comes from private sector sources while other funding comes from the government. Hence, in the notation of our model, $BERD = R + S$.

- We refer to the part of BERD that is funded by private sources as “privately-funded R&D,” or “private R&D.” This is the variable R , the main dependent variable in equations (4) and (5).
- We refer to the part of BERD that is funded by the government as “government-funded R&D” or “public R&D.” This is the variable S .
- A subset of public R&D is defense-related, and we refer to it as “defense R&D.” Note that S only includes government-funded R&D conducted by private firms, and does not include R&D conducted by universities (and other non-profits) and by the government itself (e.g. in government R&D labs). Also note that we do not have data on industry-specific defense R&D in the OECD dataset, but we instead construct a predicted defense R&D as described above.¹⁴

Appendix Table A1 summarizes the variable definition and presents summary statistics.

Facts about R&D. There is wide variation in private R&D, public R&D, and defense R&D across countries, industries, and years. Consider first aggregate R&D as a percent of GDP by country (Appendix Table A2). The most R&D-intensive country is South Korea at 2.7%, followed by Sweden at 2%. The US also has a very high R&D/GDP ratio of 1.9%. At the other end of the spectrum, there are

¹⁴ The OECD also reports two other sources of funds: Other national and funds from abroad. These sources are small, with other national sources contributing to 0.01% of total R&D, and sources from abroad contributing to 2.5% of total R&D in the dataset. We add these funding sources to privately funded R&D for simplicity, but our main results are not affected by this.

Southern European countries like Greece and Portugal, with ratios approximately 0.2%. Although there appears to be a general upward trend in R&D over time, there is substantial variation across countries in growth rates, with some countries experiencing steep increases (e.g. Denmark) while others experience declines (e.g. the UK).

R&D intensity also varies widely across industries (see Appendix Table A3). The most R&D-intensive industries are generally IT (Office, accounting, and computing machinery) and telecommunications (Radio, TV, and other communications equipment), with R&D intensities of over 20%. The next most R&D intensive sectors are chemicals (including pharmaceuticals), medical/precision instruments, and transport equipment (including aerospace) with over 10% of value added devoted to R&D. By contrast there is very little formal R&D in the distributive trades (wholesale and retail), personal services, and construction.

Public R&D also varies widely across countries and over time. Table 1 shows that the US and Eastern European nations such as Poland and Slovakia have the highest share of R&D funded by the government (over 15%), whereas the share is under 2% in Switzerland and Japan. In many countries, such as the US, the UK, France, and Canada, the rate of public funding has decreased over time. Some of this is likely to be due to a shift from direct to indirect support to business R&D, such as tax breaks (see Guellec and van Pottelsberghe de la Potterie, 1999). We explicitly add controls for tax incentives in robustness checks presented below. Table 2 shows the defense share of government-funded R&D by country. Not surprisingly, the US has the highest proportion of defense-related R&D (57%), followed by Great Britain (35%), and then France (29%). In the data, we observe the defense-related part of the government's total R&D budget from the OECD MSTI.¹⁵ Ideally, we would have just the government-funded and business-conducted part of R&D, but this data does not exist over time across countries. It is likely that the two series track each other, however. Indeed, in the case of the UK and France, both series are available and the correlations of the two series are 0.85 and 0.45, respectively.

The defense share of R&D varies not just across countries, but also within country over time. This is important for the identification of our models, which include country by industry fixed effects. Figure 1 illustrates how the four largest economies in our data experienced very different developments

¹⁵ Specifically, "Total government funded R&D" is all government budget appropriations or outlays of total R&D, i.e., not just the government-funded part of R&D conducted by businesses, but also the government-funded part of R&D conducted outside of enterprises.

in their shares of defense-related and government-funded R&D to GDP ratios over time. In the United States, defense R&D spending started at a very high level in the late 1980s under Reagan (over 0.8% of GDP) and fell subsequently after the fall of the Berlin Wall in 1989. After 9/11, defense R&D spending ramped up again under the War on Terror and the wars in Afghanistan and Iraq, rising from 0.45% (in 2001) to 0.59% (in 2008) of GDP. In Germany, defense spending is at a much lower level. Like the US, Germany reduced defense spending after the Cold War, with the rise of President Gorbachev and the fall of the Berlin Wall. In 1996, however, Germany and France cofounded a military agency focusing on R&D activities, causing a pick-up in defense R&D in Germany. In contrast to the US, Germany did not ramp up defense spending after 9/11; instead, it continued to downsize its military (European Parliament, 2011). In stark contrast, Japan has an even lower level of defense R&D spending, as its constitution commits the country to pacifism. However, Japan increased its military activities in response to North Korean missile tests in the late 1990s by starting a surveillance satellite program that resulted in satellite launches in 2003, 2006, and 2007 (Hagström and Williamsson, 2009). Finally, France shows a time-pattern relatively similar to Germany: The reduction in defense spending after the end of the Cold War is visible, but in contrast to Germany, France did ramp up defense spending after 9/11.

Overall, the experiences of these four major economies with highly variable levels of defense R&D illustrate how the timing of changes in defense R&D often reflects factors that are largely exogenous to economic and technological conditions, being driven by geopolitical events that are heterogeneous across countries.¹⁶

Our instrumental variable strategy is predicated on the notion that defense R&D is an important driver of overall government-funded R&D. Figure 2 presents the series of defense R&D and public R&D by *country* (summed across industries). Clearly, in most cases the two series tend to move together: the weighted average correlation is 0.28 (standard error 0.06). The importance of defense R&D varies widely across industries: Aerospace tends to be the single most important beneficiary of defense R&D (see Appendix Figure A1). In the OECD data the first stage of our IV relies on the relationship between public R&D and predicted defense R&D. The correlation at the industry level is visually strong (see Appendix Figure A2), with a weighted average correlation of 0.32 (with standard error of 0.07). In years when

¹⁶ Spain saw a rise in military spending after 1996, when the conservative center-right party came to power and pursued a new defense policy. The policy included a large increase in the military budget, which resulted in a sharp increase in military R&D spending. The financial crisis in 2008 forced the government to significantly cut the military budget, including R&D contracts (Miralles, 2004; Barbé and Mestres, 2007).

defense R&D is high (low), overall government funded R&D tends to be high (low). Below we quantify this relationship more formally.

3.2 French Firm-level Data

Data Sources. We use firm-level data collected by the French Ministry of National Education, Higher Education and Research (“Ministry of Research”) in their annual R&D survey from 1980 through 2015. The Data Appendix provides details on the survey that seeks to include all large firms that perform R&D and a rotating sample of smaller firms that perform R&D. Firms conducting research and development are asked to report detailed information on their R&D activity (R&D budget, number of R&D employees, R&D wage bill, number of researchers), and broader firm information (such as their firm identifier - SIREN, total number of employees, sales, main industry, etc.). Importantly for us, the dataset includes information on which firms receive R&D subsidies and how much, as well as what type of subsidy they receive (including defense). In the rest of the paper:

- We refer to all R&D subsidies originating from Ministry of the Armed Forces and its agencies as “defense R&D subsidies.”
- We refer to the sum of all R&D subsidies (including defense R&D subsidies) originating from any ministry or government agency as “total R&D subsidies” or just “R&D subsidies.”
- We refer to the firm’s R&D budget less total R&D subsidies, other national funds, and international funds, as “privately funded R&D.”

Descriptive Statistics. The sample includes 40,787 firms for an average of 3.9 years each. Only 24% of firms appear in five or more years and almost 40% of firms appear just once. Since our models include firm fixed effects, we drop firms that appear only once. We also drop firms with missing R&D data. The usable sample includes 12,539 firms appearing an average of 6.5 years each, 56% of which appear in more than 5 years. Summary statistics are in Panel B of Appendix Table A1.

The largest two-digit industries by firm count in our sample are business services (8,025 firms) and software/data (7,507 firms), followed by machinery (2,755 firms). Because the survey targets firms that are likely to be conducting R&D, this ordering also holds when we rank industries by number of firms conducting R&D in one or more years: 7,112, 6,940, and 2,550 firms in business services, software/data, and machinery, respectively. Looking at the number of firms receiving R&D subsidies,

business services (3,118 firms), software/data (3,092 firms), and machinery (1,178 firms) remain the three most prominent industries. R&D subsidies cover a large fraction of firms in these industries, while a much smaller portion receive defense R&D subsidies: just 229 business services firms, 133 electronics firms, and 115 software/data firms over the 36-year period for which we have data.

While most R&D subsidies are not for defense, most euros allocated via subsidies *are* for defense. Of the €833 billion in R&D conducted in our sample, €87 billion was publicly funded, and €57 billion of that was targeted at defense. In industries like aerospace/transport,¹⁷ the dominance of defense subsidies is even clearer, with the industry conducting €119 billion in R&D, of which €38 billion was publicly funded, almost €31 billion specifically for defense. Using the value of defense subsidies allocated to each industry, the largest defense industries after aerospace/transport are electronics, technical instruments, machinery, and chemicals.

4 The Effect of Government-Funded R&D on Privately-Funded R&D, Employment and Wages

We begin our empirical analysis by examining a case study of 9/11 (sub-section 4.1). We then estimate the effect of publicly financed R&D on privately financed R&D (sub-section 4.2) and jobs and wages (sub-section 4.3). In sub-section 4.4, we estimate international spillover effects of R&D. In section 5 we turn our attention to the effects of R&D on productivity.

4.1 An Example: the Aftermath of the 9/11 Terrorist Attacks

As we showed in Figure 1, the 9/11 terrorist attacks induced the Bush administration to suddenly increase military R&D spending in the US. We are interested in what happened to private R&D following 9/11. Figure 3 shows the differential change in private R&D intensity experienced by two “defense sensitive” industries—namely aerospace and ICT—compared to the change experienced by industries that are less dependent on defense R&D. We estimate difference in difference models using data from 1998 to 2005, with $\ln(\text{private R\&D/output})$ as the dependent variable and the defense-sensitive dummy interacted with year dummies pre- and post-2001 (as well as industry and time dummies). The figure shows the coefficients on the defense sensitive dummy in each year, together with a 95% confidence interval. It shows that before 9/11 there is no obvious differential trend in private R&D intensity. However, after 9/11, the data show a rapid increase in private R&D intensity in the defense-

¹⁷ In addition to aerospace, this includes rail and ships, but not automobiles.

sensitive sector compared to other sectors. The effect of 9/11 appears both statistically significant and economically sizable.

In Figure 4, we plot the growth in industry-specific public defense R&D intensity in the post-9/11 period compared to the pre-9/11 period on the x-axis and the growth in private R&D intensity on the y-axis.¹⁸ The figure shows a strong positive correlation (0.66, significant at the 1% level) between the industries that had the fastest increases in defense spending (like aerospace) and those that had the fastest increases in private R&D spending.

This evidence is consistent with a crowd-in effect, whereby an increase in public R&D results in additional increase in private R&D. But of course, the findings in Figures 3 and 4 need to be interpreted only as suggestive, because they could be explained by other changes in the economy associated with 9/11 that affect private firm investment in R&D. As we discussed in Section 2.3 above, the increase in private R&D could reflect an increase in future expected product demand for defense sensitive industries. It could also reflect other policy changes put in place after 9/11. To deal with these concerns, we now turn to a systematic analysis of the effect the effect of public R&D and defense R&D based on all countries and years using our OECD data; followed by a firm-level analysis using the French data.

4.2 Effect of Public R&D on Domestic Private R&D

4.2.1 Estimates Based on OECD Data

Table 3 presents estimates of the relationship between privately funded R&D and lagged public R&D in the OECD industry-country panel. The dependent variable is R&D conducted in the private sector (BERD) that is also financed by the private sector (recall that it excludes government financed R&D). As discussed in the Data section, “Public R&D” is government-financed R&D performed by private firms. All columns control for a full set of country by industry fixed effects and a full set of industry by year dummies. Standard errors are two-way clustered at the industry-country and country-year level. All models are weighted by the industry-country pair’s initial share of employment in total country employment.

Panel A of Table 3 reports OLS estimates. It shows that there is a statistically significant positive correlation between public R&D and private R&D, more consistent with crowding in than crowding out.

¹⁸ We use 1999-2000 as the pre-policy period and 2004-2005 as post policy period after the start of the Iraq War, but the exact choice of year makes little difference.

In Panel B, we report 2SLS estimates obtained by using predicted defense R&D as an instrument for public R&D. The first stages of our instrumental variable estimates are generally well identified. Weak instrument diagnostics are reported at the bottom and show that the instruments have good power: the F-Test (Kleibergen-Paap) ranges from 10 to 15 in our main specifications; and Anderson-Rubin Wald test rejects the null hypothesis of weak instruments at the 5% level, when controlling for output and GDP. The IV tests weaken slightly in the last two columns when we include (insignificant) additional variables (such as non-business public R&D and corporate tax revenues), but remain reasonable – e.g. the Anderson-Rubin Wald test still rejects the null hypothesis of weak instruments at the 10% level.

Full first stage coefficients are reported in Appendix Table A4; these are interesting in their own right. A priori, it is unclear whether an increase in predicted defense R&D in an industry will necessarily result in an increase in total government-funded R&D in that industry. In theory, given a budget constraint, it is possible that increases in defense spending are offset by declines in non-defense subsidies, leading to no net effect on total public R&D. The estimates in the table, however, suggest that this is not the case. A 10% increase in predicted defense R&D is associated with a 1.1% to 1.7% increase in total government-funded R&D, so there is not complete offset.¹⁹

The entry in column (1) in panel B of Table 3 indicates a positive effect of public R&D on private R&D. A 10% increase in public R&D subsidies is associated with an 8.1% increase in the industry's privately funded R&D spending in the following year. A comparison with panel A indicates that the point estimate is larger than the corresponding OLS estimate. This could indicate that subsidies are compensatory – targeted at “losers” and/or the presence of measurement error in private R&D.

One concern is that changes in defense R&D that are due to changes in the political orientation of a government might be correlated with changes in other policies that affect firms' private R&D spending. For example, our model would be biased if, say, right-wing governments tend to both increase defense spending for specific sectors and simultaneously adopt pro-business policies for those sectors. In columns (3) through (6), we probe the robustness of our estimates to additional controls intended to capture variation in public policies. Since these additional controls are not always available, and our sample size declines from 4,951 to 4,181, in column (2) we replicate the model in column (1) using the smaller sample for comparison. The estimated coefficient falls slightly to 0.511 (0.207).

¹⁹ There is also likely to be measurement error in the instrument that could attenuate the relationship between public R&D and defense R&D.

In column (3) of Table 3 we add controls for industry output and national GDP; the coefficient (standard error) on public R&D falls further to 0.434 (0.179). In column (4), we add a measure of R&D tax credits based on data from Thomson (2012). R&D tax credits are an alternative form of government support for R&D used by a number of countries. Over the past 20 years, many governments have started to replace direct R&D subsidies with other fiscal policies such as R&D tax credits (Guellec and van Pottelsberghe de la Potterie, 1999; Moretti and Wilson, 2012). From the point of view of governments, publicly funded R&D and R&D tax credits are likely to be substitutes, making it possible that in practice the two types of public support are negatively correlated. In this case, our estimates might understate the true effect of government-funded R&D. In practice, the magnitude of this bias is unlikely to be large, since R&D tax credits are in most countries part of the national tax code, and unlike the direct R&D subsidies, they are not industry-specific. Empirically, the coefficient on public R&D in column (4) appears to decrease only slightly, to 0.389 (0.177).

Besides businesses, other institutions like universities and government-funded research labs receive subsidies for R&D, which might be correlated with business R&D subsidies. In column (5) of Table 3, we also include R&D subsidies to non-business institutions. Empirically, non-business R&D does not appear to affect private R&D undertaken by businesses significantly, and the coefficient on public R&D rises slightly. R&D subsidies might also be correlated with other business favoring policies, for example taxes on businesses, which might also affect private R&D directly (e.g. Akcigit et al, 2018). In column (6) we control for business tax revenues as a proportion of GDP (tax revenue data is from OECD and includes taxes on income, profits and capital gains of corporates). The point estimate on public R&D subsidies is robust to this addition.

As an additional robustness check on the role played by government policies we estimate a model that controls for the political orientation of the government. If R&D subsidies are correlated with other government policies, this is likely to be especially true for the case when defense R&D changes are due to changes in the political orientation of the government after elections. Columns (1) and (2) in Appendix Table A5 indicate that the estimated effect of public R&D appears unchanged, indicating that variation in industry-specific defense R&D is not highly correlated with the general political orientation of governments.²⁰

²⁰ The political orientation data is from the World Bank's Database of Political Institutions (DPI) and indicates whether the chief executive's party is right wing, center, or left wing.

Our identification strategy—using defense R&D as the exogenous component of public R&D—is predicated on the idea that defense spending is primarily driven by geopolitical shocks and is uncorrelated with unobserved industry-specific shocks to determinants of private R&D. An important concern is that increases in defense R&D spending might be correlated with increases in expected future demand for output. For example, after 9/11, US firms producing aircraft may have anticipated increased demand for military planes and increased private R&D in expectation of this greater demand, even in the absence of public R&D. This would violate the IV strategy as both public and private R&D respond to an exogenous event. While some specifications in Table 3 conditioned on industry and aggregate output and industry by time dummies, these variables may not fully account for expectations of future demand changes that are specific to an industry-country pair.

We seek to address this concern in four ways. As a first pass, in columns (3) and (4) of Appendix Table A5, we show robustness to conditioning on *future* industry output.²¹ These models are not our preferred specification, because future output needs to be thought of as an endogenous variable. As we discussed in Section 2.2, the very reason why firms engage in investment in private R&D is indeed to increase future output. Second, in columns (5) and (6) of Appendix Table A5, we looked at using *only* country*year variation in the defense instrument. These estimates ignore variation across industries and are identified only by variation across countries and time in defense expenditures. While the first stage is a bit lower, the estimated coefficients are not very much affected, and slightly larger, if anything. Thirdly, in columns (7) and (8) of Appendix Table A5, we estimated models that condition on non-R&D military spending and expectations thereof. The sample size drops to 2,106 since we do not have data on total public military spending for all countries but the estimates based on the sub-sample where we do have this information are similar to the baseline estimates in Table 3. Including total military spending does not significantly affect the coefficients on public R&D.²²

Finally, in Appendix Table A6, we perform placebo tests based on components of defense spending that are unrelated to R&D subsidies paid to businesses: defense procurement excluding R&D; and military wage bill excluding R&D (either using a narrow or a broad definition of R&D, so overall

²¹ We also constructed the expectation of demand by running VARs of industry output against third order distributed lags of public R&D, output, and GDP. The IV estimation results in a coefficient of 0.241 (0.135), which is smaller, but still large and significant.

²² The direct effect of total military spending or procurement on private R&D is positive, as expected, but the first stage F-statistic is low.

there are four placebo instruments). Changes in both defense procurement excluding R&D and military wage bill excluding R&D are likely to be correlated with changes in demand, but should not result in changes in R&D. Therefore, in models where public R&D or private R&D is regressed on defense procurement or military wage bill, the four placebo instruments should not be predictive of public or private R&D. Finding a significant correlation between the placebo instruments and public or private R&D would suggest that our IV estimates might be driven by demand effects coming from defense spending other than R&D, or by a correlation of defense spending with other policies that encourage economic growth and therefore R&D. The results in Appendix Table A6 indicate that measures of non-R&D defense procurement and non-R&D military wage bill are uncorrelated with public and private R&D.

Overall, while we cannot completely rule out that future demand expectations play a role in our estimates, the weight of the evidence appears to be more consistent with the effects of public R&D on private R&D reflecting forces of supply rather than demand.

We performed numerous additional robustness tests. For example, since our models are identified by changes in public R&D over time, one might be concerned that the results are driven by countries with very low defense R&D levels. If we re-estimate our model only including countries with above-median defense R&D to GDP ratios—US, France, UK, South Korea, Sweden, Spain, Germany, Slovakia, Italy—our estimates appear robust. The OLS and IV coefficients are 0.209 (0.043) and 0.606 (0.210), respectively.²³ We also checked for outliers by (i) winsorizing observations in the top and bottom 1% of the *level* of R&D/output distribution and (ii) winsorizing observations in the top and bottom 1% of the *changes* in R&D/output distribution. Additionally, we tried winsorizing the growth rates of the IV. The results are robust to all these modifications.²⁴

Finally, we note that our baseline model omits country by year dummies. The reason is that the first stage is weak if country-year fixed effects are included, reflecting the fact that much of the variation in the instrument is at the country-year level. The addition of country-year dummies makes little

²³ One may also harbor the opposite concern, namely that our results could be driven by the US. The results are similar when we drop the US. The coefficient on public R&D in the IV specification is 0.489 (0.185).

²⁴ One other concern that we tested is whether the positive effect of public R&D subsidies on private R&D subsidies in the same industry might be driven by R&D subsidies to industries which are connected by input output linkages. In order to test this, we control for domestic R&D in other industries, which we weigh by their input or output share to the respective industry. This concern does not seem to be relevant in the data, as our estimates remain unchanged (results available upon request).

difference to the OLS results, especially once we include GDP to account for macroeconomic shocks.²⁵ We next turn to the richer micro data from France, where we can fully control for country by time shocks (and many other factors) even in our IV specifications.

4.2.2 Estimates Based on French Data

Table 4 contains the estimates for the French dataset. Compared to the estimates based on the OECD data, the firm-level French data allow for a much finer level of detail, since we observe which firms within an industry actually receive public R&D and which do not. In terms of identification, firm-level data allow us to estimate models that include firm fixed effects, therefore accounting for all time-invariant heterogeneity across firms. Identification comes from comparing the level of private R&D in the same firm observed before it receives a government R&D subsidy and after it receives a government R&D subsidy.

Panel A of Table 4 presents the industry-level results for France and panel B presents the firm-level results. We present industry-level results for comparison to the OECD industry-country data in Table 3, although it should be noted that the French data allow for a finer degree of industry disaggregation (195 sectors). Column (1) of panel A shows the OLS estimates. The coefficient suggests a positive correlation between privately funded business R&D and lagged government subsidies, but is smaller in magnitude than the OECD results in Table 3. Column (2) reports the corresponding IV estimate using defense spending predicted from more aggregate industry trends as an instrument for defense R&D subsidies. The first stage F-statistic is $F = 11.56$ (the first stage coefficients are reported in Appendix Table A7). The IV estimate is significant and much larger than the OLS estimate, just like the OECD results. The IV coefficient of 0.346 is not significantly different from the comparable OECD coefficient of 0.511 in column (2) of Table 3 Panel B (p-value of difference = 0.20).

Recall that defense spending at the industry level is not available in OECD industry data for most countries, but we do have it in France. Consequently, we include it directly on the right hand side of the private R&D equation in the columns (3) and (4) of Table 4. The coefficient on defense subsidies is positive and significant for the OLS and IV specifications, although again the IV coefficient is larger: 0.150 (0.041). Note that a 10% increase in total subsidies is obviously a larger amount of money than a

²⁵ For example, in the final column of Panel A in Table 3, the coefficient (standard error) on public R&D increases to 0.167 (0.029) from 0.160 (0.035) when country by year fixed effects are included.

10% increase in defense subsidies alone, which explains the smaller elasticity in column (4) compared to (2).

The firm-level analysis in panel B of Table 4 is based on a longitudinal sample of 12,539 firms observed for several years, for a total sample size of 82,015 firm-years. Panel B shows similar patterns to the results in panel A. In column (2), the IV coefficient is 0.119 (0.069), while in column (4), it is 0.374 (0.215).

The IV estimates in panel B of Table 4 again lead us to reject the null of crowd-out: increases in public R&D result in more investment in private R&D, not less. Based on entries in column (2), a 10% increase in R&D subsidies is associated with a 1.2% increase in the firm's privately funded R&D spending in the following year. This confirms that even after controlling for firm fixed effects, defense R&D subsidies appear to be crowding in private R&D spending.

A comparison with panel A of Table 4 indicates that industry-level coefficients are smaller than firm-level coefficients when we use defense R&D subsidies (columns (3) and (4)), but the reverse is true for total R&D subsidies (columns (1) and (2)). Coefficients from industry-level data do not need to be identical to coefficients from firm-level data in the presence of technology spillovers from R&D within an industry. Industry coefficients should be larger if crowd-in induces rival firms to do more R&D (strategic complementarity). However, it might be that rivals do less R&D if there is strategic substitutability (e.g. free riding), for example, meaning that industry coefficients would be less than their firm-level counterparts (Bloom, Schankerman, and Van Reenen, 2013). We will investigate spillover effects at the international level in more detail below.²⁶

Overall, there is little evidence in Tables 3 and 4 of upward bias in the OLS estimates. In fact, the OLS estimates are consistently below the IV estimates. In the context of our discussion in Section 2, this is consistent with compensatory government policies, whereby governments tend to subsidize industries that are underperforming in terms on R&D investment.²⁷ The finding of IV estimates larger than OLS estimates may also reflect attenuation bias from measurement error in OLS; or a local average treatment effect, with public funds directed towards industries that are likely to “match” subsidies more strongly (or a combination of these explanations).

²⁶ When we separate the sample into firms that are larger than the median in their industry or smaller than the median in their industry, OLS estimates indicate that that smaller firms crowd in less than large firms, though these estimates are imprecise and preclude definitive conclusions.

²⁷ Criscuolo et al. (2019) find that the same appears to be true in the case of UK investment subsidies.

4.2.3 Magnitude of the Estimated Effect

Taken together, the estimates in Tables 3 and 4 indicate that increases in public R&D translate into increases in private R&D expenditures. This is true both when we focus on industry changes across the whole OECD or within France and when we focus on within-firm changes in France. This crowd-in is consistent with the existence of agglomeration economies—whereby increases in government R&D raise the returns for private companies in the same country and industry—or large fixed costs or credit constraints.

Our preferred elasticity for the OECD dataset is 0.434 (Table 3, panel B, column (3)), suggesting that a 10% increase in government subsidies in a given year is expected to result in a 4% increase in private sector R&D the following year. This implies that \$1 of additional public funds for R&D translates into \$5 of extra R&D funded by the private sector at the mean values of public and private R&D. Our preferred elasticity at the firm level in the French dataset is 0.119 (Table 4, panel B, column (2)). This implies that €1 of additional public funds for R&D translates into €0.85 of extra R&D funded by the private sector at the mean values of public and private R&D in France. As noted above, the smaller firm level effects compared to industry level effects could be due to positive within industry spillovers.

Our findings imply that in some industries, defense-related R&D is responsible for a significant portion of private R&D investment. For example, in the US “aerospace products and parts” industry, defense-related R&D amounted to \$3,026 million in 2002 (nominal). Our estimates suggest that this public investment results in \$1,632 million of additional investment in private R&D.²⁸ If we take the total amount spent by the US government, we estimate that private R&D investment is \$71 billion higher than the counterfactual with no government-funded defense R&D.

Interestingly, differences in defense-related R&D can account for some of the differences in private R&D across countries. For example, our estimates indicate that if France increased its defense R&D to the level of the US as a fraction of GDP (admittedly a very large increase: roughly a factor of 2.2), private R&D investment would increase by 8.7%.²⁹

²⁸ Public R&D in aerospace products and parts amounted to \$4,306 million in 2002 (nominal). Therefore, defense related R&D was 70% of public R&D. A reduction of public R&D by 70% translates into a reduction of private R&D by 30% using our preferred elasticity of 0.434 in column (3) of Table 3. Private R&D was \$5,349 million in aerospace products and parts in 2002, so 30% of this number yields \$1,632 million.

²⁹ Our estimates also indicate that if Germany increased its defense R&D to the level of the US as a fraction of GDP (an even larger increase), private R&D investment would increase by 60%. Defense R&D expenditure as share of GDP in our data

4.3 Effect on Employment and Wages

We now turn to the effect of public R&D on employment and wages. The discussion in Section 2 suggests that if the supply of R&D workers is inelastic in the short run, increased R&D spending could translate into significantly higher wages (Van Reenen, 1996; Kline et al., 2019), especially of R&D workers (Goolsbee, 1998). Thus, it is possible in principle that finding increases in private R&D expenditures following increases in government-funded R&D may simply reflect an increase in the cost of inputs used in R&D, namely specialized labor.

Table 5 reports IV estimates of models similar to those in Tables 3 and 4 where the dependent variable is employment or wages. We distinguish between R&D personnel and non-R&D personnel. (The corresponding OLS estimates are reported in Appendix Table A8.) For convenience, column (1) reproduces the baseline effect of public R&D on private R&D from column (1) of Table 3 (OECD data) and column (2) of Table 4 (French industry and firm data). Since data on employment and wages are not available for all countries and industries and years, column (2) reports the baseline elasticity estimated on the sample for which employment and wage data are available.

In column (3) of Table 5, we focus on workers directly engaged in R&D activities. We uncover positive and significant effects. The coefficients for OECD industry data is 0.563 (0.329), while the corresponding coefficients for French data are 0.394 (0.069) at the industry level and 0.114 (0.049) at the firm level. Notably, the estimated elasticities are similar to those for R&D expenditures in column (2), indicating that employment of scientists increases proportionally to increases in R&D expenditures. The large employment effect is consistent with an elastic labor supply, indicating that it is easy for R&D workers to relocate to the affected industry from other industries or countries.

The estimates in column (4) of panels A and B of Table 5 uncover smaller effects on employment for non-R&D workers. The elasticity for non-R&D employment in OECD data is only 0.127 (0.048), while in French industry data it is 0.324 (0.086). This result is consistent with shared fixed costs, whereby the marginal R&D activity induced by public R&D requires new scientists and non-scientific personnel but utilizes existing labs. It is also consistent with the notion that the effect of labor-augmenting

(averaged over all years) is 0.58% for the US, 0.28% for France, and 0.07% for Germany. To reach the defense R&D share of the US, Germany would have to increase own defense R&D spending by a factor of 9.3—a very large increase. Multiplying this with the coefficients of the first stage and the IV estimates yields the resulting percent increase in private R&D spending.

innovations is somewhat larger in magnitude than the labor-saving innovations.³⁰ In the French firm-level data the coefficient is positive but statistically indistinguishable from zero.

The dependent variable in column (5) of Table 5 is the average salary of R&D workers.³¹ The key coefficient in all panels is positive but imprecisely estimated and indistinguishable from zero. The lack of clear wage gains following increases in public R&D strengthens the case for a highly elastic supply of R&D personnel to the industry.

4.4 International R&D Spillover Effects

So far, we have estimated the direct effect of government-funded R&D in an industry and country on private R&D activity in the same industry and country. We now consider the possibility that government-funded R&D in an industry and country may have an additional, indirect effect on private R&D investment in *other* countries. It is possible that increases in government-funded R&D in an industry in a given country will result in lower R&D in similar industries abroad. Alternatively, knowledge generated from public R&D in one country may spill over to a different country and generate additional private research efforts. To quantify these spillover effects, we estimate equation (6) where we include both domestic public R&D and neighbors' public R&D. We base this analysis on the OECD industry-country data.

Specifically, in panel A of Table 6 we regress private R&D on lagged domestic public R&D and lagged neighbors' public R&D, measured as a weighted average of public R&D in other countries in the same industry and year, with weights reflecting various measures of geographic and economic proximity. Column (2) uses the difference in GDP per capita as a distance measure, column (3) the geographic distance in kilometers between the capital cities, column (4) the difference in skill intensity as measured by the share of the population with tertiary education, column (5) the similarity of patent technology classes (out of 15 different technology classes), column (6) the FDI flows that a country receives from other countries, column (7) the difference in R&D intensity as measured by R&D/GDP. We instrument

³⁰ We added the output terms to examine this and the effect of public R&D is somewhat smaller with an elasticity (standard error) of 0.080 (0.028).

³¹ We measure average salary of researchers by dividing the costs of R&D personnel by the number of R&D personnel. For the OECD data, we can also estimate the effect of public R&D on overall average salaries but find no effect (we measure average salary by dividing total wage bill by total employment).

domestic public R&D but not neighbors' public R&D, which we assume to be exogenous to domestic firms. (The corresponding OLS coefficients are in Appendix Table A9).

The effect of domestic public R&D on private R&D remains positive across all specifications. The coefficient of interest is the one on neighbors' public R&D. It is positive and statistically significant in all columns but (4) which uses skill intensity, indicating that if one country increases its public R&D, firms in nearby countries in the same industry increase their investment in private R&D after controlling for public R&D received from their own government. The positive effect is consistent with significant crowd-in between close countries when proximity is defined by income, geography, technology, FDI flows, and R&D intensity. The estimated elasticities range from 0.158 (0.093) in column (2)—where proximity is based on per-capita GDP—to 1.014 (0.281) in column (5)—where proximity is based on technology.

Panel B of Table 6 uses business R&D (not government-funded R&D) to compute the international spillover pool to test whether private R&D undertaken in a foreign country can also generate spillover effects. Here five of the six coefficients are negative (two of them significantly so), with elasticities that are somewhat smaller than those in panel A, providing evidence of international displacement between firms. Put differently, private R&D investment by firms in an industry and country appear to have no effect, or discourage competitors in the same industry in other countries from undertaking their own R&D.

In panel C of Table 6, we include both international public R&D and international business R&D simultaneously. In all but one specification, elasticities on international public R&D are positive while elasticity on international private R&D are negative and significant. The elasticities on the international public R&D and international private R&D appear of the same order of magnitude in absolute value, indicating that they have a quantitatively similar percentage effect, although of opposite sign.

A comparison between panels A and B indicates that private and public R&D generate different spillover effects on R&D behavior. While there are positive spillover effects from public R&D subsidies, private R&D leads to either no spillover effects or crowding out. One possible explanation could be the fact that governmental subsidies may be associated with requirements to make research findings public; or that research supported by the government tends to be more basic research.

5 The Effect of R&D on Productivity

Having documented the effect of government-funded R&D on private R&D investment, we now turn our attention to quantifying the effect of R&D investment on productivity. We start in sub-section 5.1 by deriving the econometric model that we use in our empirical analysis. In sub-section 5.2 we present estimates of the effect of total R&D on domestic productivity based on industry-level OECD data and firm level French data. We also present estimates of international spillovers based on OECD data—namely, the effect of R&D performed in an industry and country on productivity in the same industry but in different countries. We conclude in sub-section 5.3 with an illustration of the magnitude of the effects implied by our estimates based on the 9/11 shock.

5.1 Econometric Models

We relate measures of productivity to R&D investment. In the OECD data, we measure productivity as industry-country-year TFP. We approximate $F(\cdot)$ in equation (1) above by a second order flexible form and can therefore estimate TFP as the superlative index:

$$\Delta \ln A_{ikt} = \ln \left(\frac{VA_{ikt}}{VA_{ik(t-1)}} \right) - \frac{1}{2} (\theta_{ikt} + \theta_{ik(t-1)}) \ln \left(\frac{L_{ikt}}{L_{ik(t-1)}} \right) - \left(1 - \frac{1}{2} (\theta_{ikt} + \theta_{ik(t-1)}) \right) \ln \left(\frac{K_{ikt}}{K_{ik(t-1)}} \right) \quad (7)$$

where VA_{ikt} is value added, L_{ikt} is total employment, and K_{ikt} is the capital stock. We measure θ using labor's share of value added in two ways. First, we simply use the industry-specific average of the share (across all countries and years).³² Alternatively, we use Harrigan (1997) smoothing methods to construct the share of labor in value added, θ_{ikt} . Both methods give very similar results. The Data Appendix describes in more detail how we estimate TFP. To estimate the effect of R&D on TFP at the industry level, we assume that the total R&D stock can be described by the perpetual inventory formula: $G_{ikt} = R_{ikt} + S_{ikt} + (1 - \delta)G_{ik(t-1)}$ where δ is the depreciation rate of knowledge. If δ is close to zero, the TFP growth equation can be approximated by:

$$\Delta \ln A_{ikt} = \rho \left(\frac{R+S}{VA} \right)_{ik(t-1)} + \gamma \Delta X_{ikt} + \Delta u_{ikt} \quad (8)$$

where $\rho = \frac{\partial Y}{\partial G}$ is the gross rate of return to R&D capital.³³

³² Since our TFP estimates are at the industry level, this method is more suited to estimate productivity than alternative methods such as Olley-Pakes (1996) or Levinsohn-Petrin (2003) which are designed for estimation at the level of the firm.

³³ Although conventional, assuming $G_{ikt} = R_{ikt} + S_{ikt} + (1 - \delta)G_{ik(t-1)}$ is restrictive. We considered an alternative specification $G_{ikt} = R_{ikt} + \mu S_{ikt} + (1 - \delta)G_{ik(t-1)}$ that allows a dollar of public R&D to have a different effect on the

In the French firm-level data, we have no information on capital stock or value added. Thus, our dependent variable is labor productivity defined as output per worker, $\left(\frac{Y}{L}\right)_{fikt}$:

$$\Delta \ln \left(\frac{Y}{L}\right)_{fikt} = \rho \left(\frac{R+S}{Y}\right)_{fik(t-1)} + \Delta u_{fikt} \quad (9)$$

Equations (8) and (9) focus on the direct effect of a country's R&D investment on its own productivity. It is possible that there is an additional *indirect* effect in the form of an international technological spillover. This would occur if a country's investment in a given industry ends up benefitting the productivity of firms in different countries due to international knowledge spillovers. To test for this possibility, we use our OECD data to estimate a variant of equation (8):

$$\Delta \ln A_{ikt} = \rho((R+S)/VA)_{ik(t-1)} + \kappa(RP/VA)_{ik(t-1)} + \gamma \Delta X_{ikt} + \Delta u_{ikt} \quad (10)$$

where $(RP/VA)_{ik(t-1)}$ is the weighted average of R&D/value added in all other countries in the same industry and year, with weights measuring the economic or geographic distance of country k to all other countries (as in the spillover variable in equation (6)). A positive (negative) κ would indicate the existence of positive (negative) international technological spillovers.

In practice, variation in value added and output per worker reflect both variation in physical productivity as well as variation in the prices of output. This is a common problem in the estimation of production functions. In our context, this problem is likely to be serious because shocks to the demand for defense products (geopolitical shocks, leadership changes, etc.) are likely to result in shocks to the price of defense-related products. The defense industry is highly concentrated and has significant barriers to entry, at least in the short run. This means that the supply curve is almost certainly not infinitely elastic in the short run. An upward sloping supply curve implies that when product demand increases, our measure of TFP increases even if productivity does not change. As standard in this literature, we deal with this problem by using industry-year specific price deflators.³⁴

knowledge stock than a dollar of private R&D. This implies including two separate R&D terms on the right hand side of equation (8). The problem is that this requires an additional instrument for privately funded R&D. We considered using R&D tax credits, but the first stage had insufficient power when both public and private R&D were taken as endogenous.

³⁴ This is more problematic for the firm level results when there are firm-specific prices. In this case the coefficient on R&D includes an effect on markups as well as quantity-based productivity.

5.2 Effect of R&D on Productivity

In Table 7, we estimate equations (8) and (9) by regressing changes in productivity on lagged total R&D intensity. Total R&D is measured as the sum of private R&D (R), public R&D (S) and R&D from any other source. Columns (1) to (4) are based on OECD industry-level data (equation (8)). Columns (5) to (8) are based on the French firm-level data (equation (9)).

The OLS coefficient on R&D of column (1) indicates a positive correlation between lagged R&D intensity and subsequent TFP growth. Column (2) adds country dummies and shows that the coefficient is essentially the same as in column (1). We find a slightly smaller coefficient of 0.064 (0.027) for the IV estimate in column (3), but this effect is not statistically different from the OLS estimate. Column (4) reports the reduced form estimate. In this model the independent variable is defense R&D divided by value added, which has a both positive and statistically significant effect.³⁵ We conduct a similar exercise for French three-digit industries in columns (5) and (6). The OLS estimates are positive, just like those for the OECD industry-country panel, but of smaller magnitude—0.026 (0.004) in column (5). Estimates are unchanged when we add two-digit industry fixed effects in column (6). In columns (7) and (8) we repeat this exercise on the French firm-level data. We uncover coefficients (standard errors) of 0.040 (0.004) and 0.043 (0.004), larger than the corresponding industry-level elasticities.

Unfortunately, the IV estimates corresponding to columns (5) to (8) in Table 7 are unidentified as the first stages have insufficient power (the F-statistics at the industry and firm level are 0.96 and 0.10 respectively).³⁶

The magnitude of the estimated effects in Table 7 is economically significant. Using the estimate for the OECD industry-country panel in column (2), for example, a permanent increase in the predicted defense R&D to value added ratio of one percentage point is associated with an increase in the annual growth rate of TFP of 0.05 percentage points. Since average annual TFP growth in our sample is around

³⁵ We probed the robustness of our finding using several alternative specifications. For example, estimates of the impact of R&D on TFP using Harrigan (1997) smoothing techniques are very similar: the equivalent coefficient (standard error) on R&D intensity in column (3) was 0.060 (0.026). We have also tested for heterogeneity of the effect. The effect of R&D on TFP seems slightly more pronounced in less R&D intensive countries. It also varies across industries, but this variation is not systematically associated with industry R&D intensity.

³⁶ The lack of first stage power in this table stands in contrast with strong first stages in the previous tables. A possible reason is that the first stage does not include firm fixed effects (otherwise the second stage would be effectively in double differences). This suggests the *changes* in firm R&D and *changes* in firm defense public may be more strongly related than the level relationship.

1.01%, this represents an increase from 1.01% to 1.06% a year. Using the OLS estimate in column (6), a similar calculation suggests an increase from an average annual labor productivity growth rate of 0.05% to 0.076% a year in France.

To put this in perspective, consider that our estimates indicate that if France and Germany were to raise their defense spending to the level of the US as a percentage of value added—holding constant everything else and ignoring the additional tax revenues needed—they would experience an increase in the productivity growth rate by 9% and 3%, respectively.³⁷

We also note that the finding of a positive effect on productivity has implications for our earlier analysis of crowding in. The estimated positive effects on productivity indicate that the positive crowding in effects that we found in the previous section are unlikely to be driven by a relabeling of other expenditure as R&D expenses in order to obtain subsidies. A pure relabeling would not result in these productivity increases.

Finally, we turn to the question of international productivity spillovers. We estimate equation 10, where we use the same set of weights used in Table 6. Findings in Appendix Table A10 indicate statistically significant positive international spillovers with coefficients ranging between 0.14 and 0.22 in the IV specifications. In order to interpret the magnitude note that the average values of own R&D/value added and the international spillover pool are different, so the coefficients cannot directly be compared. The effect of the spillover pool is larger because it requires all the other countries to increase their public R&D.³⁸

5.3 An Illustration of the Magnitude of the Effects

We end with an illustration of the magnitudes of various effects arising from our analysis based on a concrete example: the increase in military R&D that occurred in the US after the 9/11 attacks. As described in Figure 1, US defense R&D rose by 52% bringing it from 0.45% to 0.60% of GDP between

³⁷ The average defense R&D to value added ratios for the US, France, and Germany in our data are 0.61%, 0.37%, and 0.07%, respectively. Thus, France would need to increase its defense R&D/value added by 65%, while Germany would need a 771% increase.

³⁸ For example, in a specification using geographical distance weights for spillovers, the own R&D coefficient is 0.036 and the spillover coefficient is 0.205 (Table A10, panel B, column (3)). An increase by one standard deviation of own R&D/value added (14.4 percentage points) is associated with future TFP growth by 0.52%. By contrast, an increase by one standard deviation of the spillover pool (9.7 percentage points) increases TFP growth by 2.0%. Note that the coefficient on domestic R&D falls and becomes insignificant when spillovers are included which is due to the fact that domestic and international R&D are highly correlated, so it is difficult to estimate both effects simultaneously.

2001 and 2004. Using all the linkages in the results above, we calculate that the 9/11 shock induced an increase of the TFP growth rate by 0.006 percentage points, holding taxes constant. Since the average annual growth rate in this period is 0.3%, this effect amounts to a 1.9% increase. In addition, TFP growth in other OECD countries is estimated to rise by 0.045% on average.

We provide a detailed explanation of how we quantified these effects in Appendix B. In short, the effect on US TFP growth comes from using the TFP estimate in Table 7, column (3), combined with an increase in total US R&D/GDP of 0.09 percentage points (or 4.6%; which is composed of public R&D rising by 8.7% according to Table A4, column (3), and private R&D rising by 3.8% according to the IV results of Table 3, column (3), due to crowding in). The effect on other OECD countries comes from the international spillover arising from the increase in US R&D (Table A10). There is also a small TFP offset because US R&D displaces some foreign R&D (down -0.03% on average; Table A11).

In interpreting these estimates, two points need to be highlighted. First, our estimates quantify the TFP gains occurring within a relatively short horizon. It is in principle possible that the effects are larger when looking at a longer time horizon, e.g. over decades, therefore our estimates are likely to be a lower bound of the true effect of public R&D subsidies on private R&D and productivity growth.

Second, and crucially, these estimates do not account for the opportunity cost of public funds needed to finance the increase in public R&D. In reality, the US government faces a budget constraint, and the increase in public expenditures on R&D needs to be thought of as coming either from a decrease in other forms of public expenditures or as an increase in current or future taxes, or a combination of both. In either case, the opportunity cost of the additional public R&D expenditures induced by 9/11 is unlikely to be zero, although it is difficult to estimate in practice. The opportunity cost could be high if it cut other forms of public expenses with high returns or if the expected increase in taxes crowd out private investment. Alternatively, the opportunity costs could be low if the federal government financed the increase in public R&D by cutting unproductive or wasteful public expenses. For example, military procurement is widely considered inefficient and ripe with waste and rent seeking. Reallocating military expenditures away from inefficient uses toward funding for R&D would have limited opportunity costs.

Overall, our estimates need to be interpreted as partial equilibrium estimates of an increase in public R&D: the ultimate effect on US TFP growth inclusive of general equilibrium effects is likely to be smaller (e.g. Akcigit, Hanley and Stantcheva, 2017).

5 Conclusions

Given the central role that productivity plays in explaining economic growth, it is not surprising that the search for the determinants of productivity has been central to modern economics. While a large body of empirical research has argued that R&D is a key source of firm productivity growth (e.g. Aghion and Howitt, 1992; Romer, 1990), the question of exactly which policies governments should adopt to foster R&D investment is still largely an open question.

In this paper, we study the impact of public R&D spending—and defense R&D spending in particular—on private R&D spending and productivity. We uncover some unexpected intellectual spoils of war. First, our results suggest that government R&D “crowds in” rather than “crowds out” private R&D. Second, we find evidence that an increase in public R&D in one industry and country raises private R&D in the same industry in other countries through a positive spillover effect. Third, we uncover significant effects of private R&D investment on TFP growth and therefore economic growth.

In terms of policy implications, our estimates point to a concrete tool that governments can use to raise private R&D investment in their jurisdiction. Our estimates indicate that government-funded R&D in general—and defense R&D in particular—are effective at raising a country’s total expenditures on innovation in a given industry. The ultimate effect of government-funded R&D on overall R&D significantly exceeds its dollar value because government-funded R&D stimulates additional R&D investment on the part of the private sector. This positive effect of government-funded R&D on private R&D is important not just in itself, but because it generates higher productivity. This of course does not imply that it is efficient to raise defense R&D or government-funded R&D across the board, since government-funded R&D clearly has an opportunity cost in the form of taxpayer money used plus any welfare loss that inevitably comes from taxation. Our paper does not compare the benefits of government-funded R&D to its costs.

Our findings also indicate that the benefits of public R&D investment do not stop at a country’s borders, but spillover to other countries. This implies that countries that spend aggressively on government funded R&D—like the US—indirectly support the productivity of countries with less government-funded R&D. This externality indicates the desirability of more international cooperation in government-funded R&D.

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APPENDIX A: DATA

(A1) OECD DATA

Main Variables. R&D data was obtained from the OECD database Main Science and Technology Indicators (MSTI). This dataset contains industry level R&D conducted by businesses (“Business Enterprise R&D” or “BERD”) for 26 countries between 1987 and 2009. This is our main R&D variable, but note that BERD is the main component of general R&D (“GERD”) which also includes R&D conducted by non-businesses such as universities and government R&D labs (it is not possible to break this down by industry).

The panel is unbalanced as some data is not available for some countries especially in the early years. Note that BERD is R&D conducted by business but can be financed from several sources. The OECD breaks BERD into three sources of funding: government, business and abroad. Our main variable is BERD that we generally refer to simply as “R&D”. The variable “Public R&D” is the government funded part of BERD (S). We used the OECD’s PPP US\$ values to deflate all nominal variable such as GDP, output and R&D. From the same dataset, we also obtained the total number of R&D personnel by industry and country (all measured in full-time equivalent on R&D activities) which is broken down into scientists (“# researchers”) and other R&D personnel such as lab technicians. We also have the total wage bill of these R&D personnel that enables a crude construction of the “R&D wage” of R&D labor cost divided by total R&D personnel.

For the main regressions where we use R&D as the dependent variable, we use the industry-funded part of BERD as the dependent variable. We refer to this as “private R&D” (R). For the regressions where we use TFP growth as the dependent variable, we use R&D as the key right hand side variable - i.e. total BERD from all sources of funds. We take output, employment, wage bill, capital and value added data from the OECD STAN database. We use the level of industry aggregation that maximized the matches between the databases STAN and MSTI, ending up with 26 industries (see below).

Data for missing years was linearly interpolated for all variables (between the first and the last year available per country), but no data was extrapolated forward or backwards. These imputations accounted for between zero and 35% of the final sample, depending on the variable. The results are robust to dropping interpolated values.

The industry level breakdown of BERD by source of funding was missing for the United States after 2001 in the MSTI because a change in industry classification of the OECD (to ISIC Rev. 3) required a new crosswalk from the industry classification used in the US, NAICS. However, we obtained the original R&D data from the National Science Foundation, Survey of Industrial Research and Development (SIRD), and implemented the crosswalk to our data set, verifying that the totals (all source of funds) by industry matched the totals published by the MSTI. The crosswalk is available on request.

Total Factor Productivity, TFP. TFP growth can be measured by a superlative index derived from the translog production function (as in Caves et al, 1982). This results in the following expression for industry-country-year TFP:

$$\Delta \ln A_{ikt} = \ln \left(\frac{VA_{ikt}}{VA_{ikt-1}} \right) - \frac{1}{2} (\theta_{ikt} + \theta_{ikt-1}) \ln \left(\frac{L_{ikt}}{L_{ikt-1}} \right) - \left(1 - \frac{1}{2} (\theta_{ikt} + \theta_{ikt-1}) \right) \ln \left(\frac{K_{ikt}}{K_{ikt-1}} \right)$$

where real VA_{ikt} is value added, L_{ikt} is total employment, K_{ikt} is the capital stock, and θ_{ikt} is the share of labor in value added.

Value added is from the STAN database, we use the variable VALK that gives value added in volumes. STAN uses volume indices provided by national statistical agencies that are typically derived by applying detailed deflators based on Producer Price indices (PPIs) or Consumer price indices (CPIs)

coming from detailed surveys. Volumes for activity groups are either fixed-weight Laspeyres aggregates or annually re-weighted chained aggregates of the volumes of detailed sectors.

Capital stock is given in STAN in variable CPGK, “gross capital stock (volumes)”. For some countries (e.g. US, Japan, Australia, Norway, Sweden) the capital stock was not available in the STAN database, but investment data was (variable “GFCK”, gross fixed capital formation in volumes). For these countries, we follow Scarpetta and Tressel (2002) and OECD (1999) and construct capital stock using a perpetual inventory model, which simulates the process of capital accumulation using past investment data that is adjusted for scrapping. Gross capital stock GCS_t at time t is:

$$GCS_t = \sum_j INV_{t-j} * (1 - \delta_j)$$

where INV_{t-j} is investment undertaken in period $t-j$, i.e. with vintage j at time t , and $1-\delta_j$ is the survival coefficient that represents the fraction an investment of vintage j that is still in use in year t contributes to gross capital stock in year t . The survival coefficient is between 0 and 1, and falls with vintage. We follow the literature and assume that depreciation follows a delayed linear retirement pattern, where scrapping starts only 5 years after the investment has been undertaken.

Average services lives (ASL) are usually used to measure the depreciation process of an investment. If depreciation starts after 5 year and is linear, ASL is given by

$$ASL = 5 + \frac{n}{2}$$

where n is the time period during which an investment has positive depreciation (and $n+5$ is the total lifetime of an investment). Thus, with linear depreciation, in each year depreciation equals a constant fraction δ of past investment (starting after 5 years):

$$\delta = \frac{1}{n} = \frac{1}{2 * (ASL - 5)}$$

Gross capital stock can therefore be constructed using the recursive relationship:

$$GCS_t = GCS_{t-1} + INV_t - \frac{1}{2 * (ASL - 5)} \sum_{j \geq 5} INV_{t-j}$$

Since our investment data starts only at the beginning of the sample period and we need past investment for the construction of gross capital stock, we assume that past, unobserved annual investment is equal to average investment across the sample years (separately for each industry within each country). Our results are robust to assuming past investment to be equal to initial observed investment, or average investment across the first observed years. This assumption about past investment can then also be used to calculate the gross capital stock at the beginning of the sample period. Average services lives (ASL) by country and industry are given in OECD (1999), p. 48. For countries without ASL data, we use the US industry-specific ASL.

The share of labor in value added is measured in two ways. First, we simply used the industry-specific unweighted average of the wage bill over value added (across all countries and years). Alternatively, we use the Harrigan (1997) smoothing method to construct the share of labor in value added α_{ikt} , by industry, country and year. For this, we ran a pooled OLS regression of the form

$$\theta_{ikt} = \theta_{ik} + \phi_j \ln \left(\frac{K_{ikt}}{L_{ikt}} \right)$$

where θ_{ikt} is the observed labor share in value added, α_{ik} are a full set of industry-country pair fixed effects and K and L are capital stock and employment as described above. We then used the predicted values from this regression for the industry-country-year specific share of labor in value added θ_{ikt} . As robustness checks we also used the overall average labor share in our data, or a constant labor share of 0.65 to construct TFP, but our results were not sensitive to the specification of the share of labor in value added.

(A2) French Data

We use data collected by the French Ministry of National Education, Higher Education and Research (which for brevity we will refer to as “Ministry of Research”) for their annual R&D survey from 1980 through 2015. Firms conducting research and development activity are asked to report detailed information on their R&D performance—R&D budget, number of R&D employees, R&D wage bill, number of researchers—and broader firm information—SIREN (firm ID), total number of employees, sales, main industry. The survey seeks to include all large firms that perform R&D and a rotating sample of smaller firms that perform R&D. In particular, firms with R&D expenditures above a threshold are surveyed every year, while a rotating sample of firms below the threshold is surveyed for two consecutive years. The threshold varies over time. For example, in 2007 it was €250,000 and by 2011 it was raised to €400,000. For our purposes, the net effect is that there is significant turnover in the sample that we work with among the smaller firms. We observe 40,787 firms an average of 3.9 years each. Almost 40% of firms appear just once; and only 24% appear in five or more years. Missing R&D data and our empirical strategy (firm fixed effects) reduce the usable sample to 12,539 firms appearing an average of 6.5 years each, approximately 56% of which appear in more than five years. Overall, our sample over-weights large firms that conduct significant amount of R&D.

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Defense R&D subsidies. Importantly, the survey asks about R&D subsidies granted under the authority of individual ministries, i.e. Ministry of Research or Ministry of the Armed Forces. We treat all R&D subsidies originating from Ministry of the Armed Forces and its agencies as defense R&D subsidies.

Total R&D subsidies. We define total R&D subsidies as the sum of all R&D subsidies (including defense R&D subsidies) originating from any French ministry or government agency.

Privately funded R&D. We call the firm’s R&D budget less total R&D subsidies privately funded R&D.

Industry classification. During this period French industry codes change three times (in 1993, 2003, and 2008) so we use a combination of correspondence mappings and crosswalks, created from firms surveyed under multiple classification systems, to harmonize all 4-digit industry codes to the NAF revision 1 classification in use from 2003-2008. We have 689 4-digit industries, 216 3-digit industries, and 59 4-digit industries.

In addition to the variables mentioned above, we also use information on firm sales and employment. In our data the mean (median) sales and employment for a firm-year are €260 million (€19.8 million) and 925 (129) workers. Similarly, in 2000 prices, the mean (median) firm-year receives €576,000 (€0) in defense R&D subsidies, €843,000 (€0) in total R&D subsidies, and expends €6 million (€710,000) on private R&D.

(A3) Instrument for Defense R&D Subsidies.

Basic Approach

One simple identification strategy would be to use observed defense subsidies in a year as the instrument for government funded R&D (S_{it}) in that year. However, there is a risk that even defense subsidies may have an endogenous component, responding to supply or demand shocks. Consequently, our preferred baseline IV is *predicted* defense R&D. Although the exact form of the IV will depend on the dataset and level of aggregation, the generic form of our instrument is:

$$DR_{it}^{IV} = share_{i(t-1)}^l \cdot \widetilde{def}_t \quad (A1)$$

where $share$ is a lagged industry-based share of defense R&D and \widetilde{def}_t is defense R&D at a higher level of aggregation than the endogenous variable (government funded R&D).

OECD data

Consider first the IV used in our OECD analysis. Here, the endogenous variable is S_{ikt} , government funded R&D in industry i in country k at time t . For a country k like Japan our implementation of equation (A1) is:

$$DR_{it}^{IV} = share_{i(t-1)}^{US} \cdot DefGBAORD_{kt} \quad (A2)$$

where in this case $\widetilde{def}_t = DefGBAORD_{kt}$, is the value of Government Budget Appropriations or Outlays for defense related R&D by country-year from OECD MSTI.³⁹ Note that we are not using the industry specific values of (for example) Japanese defense R&D subsidies in equation (A2), but rather the Japanese economy-wide expenditure, which reduces the risk that the cross-industry allocation in Japan could be endogenous to current industry-year shock in Japan.

The industry variation comes from the weight, $share_{i(t-1)}^{US}$, which is the share of the industry's defense R&D in economy-wide defense R&D in the *United States*. Using the US specific industry breakdown (instead of Japan's) should again mitigate concerns that the industry allocation of defense R&D could be endogenous.

French Firms

Next, consider our firm-level instrument in France. In this case our implementation of equation (A1) is

³⁹ Ideally, we would like to know the part of business enterprise R&D that was solely related to defense spending and then to break this down into the government and privately financed components, but this is unavailable. In the time series, however, it is likely that increases in $DefGBAORD_{kt}$ are strongly associated with increases the part of public R&D that goes to defense. Our first stages regressions confirm that this relationship is strong in the data. Further, in the UK, we have a more detailed breakdown and we confirm that the correlation is 0.85.

$$DR_{f,i4,t}^{IV} = share_{i4} \cdot \overline{def}_{i3,t} \quad (A3)$$

The weight, $share_{i4}$, $share_{i4,(t-1)}$, is the share of public expenditure on business conducted defense R&D in the firm's four-digit industry ("i4") as a share of the three-digit industry ("i3") that the four-digit industry is located in (aggregated up from the firm level data) averaged over the entire sample period. Our preferred method uses a time invariant average because $share_{i4,t}$ can be noisy due to the changing year by year stratified random sampling. Further, we define $\overline{def}_{i3,t} = def_{i3,t-1} + \Delta \overline{def}_{i3,t}$ where Δ is the first difference operator and $\overline{def}_{i3,t} = def_{i3,t} - def_{f,t}$ is three-digit public R&D defense subsidies leaving out firm f 's defense R&D subsidies. Therefore $\overline{def}_{i3,t}$ is close to the "leave-out mean. The reason for excluding $def_{f,t}$ from the construction of the IV is that the endogenous variable $S_{f,t}$ is partly composed of $def_{f,t}$ (the other component is the firm's non-defense related part of R&D subsidies). Thus, including $def_{f,t}$ in the IV would lead to a mechanical (and potentially endogenous) correlation between the endogenous variable we want to instrument and the instrument itself.⁴⁰

French Industries

Finally, consider the French industry level analysis. When pooled in the OECD regressions we use equation (A2). However, when we implement the more disaggregated industry analysis of (say Table 4, Panel A) we can disaggregate the IV further. In particular the implementation of equation (A1) is:

$$DR_{i3,t}^{IV} = share_{i3} \cdot \overline{def}_{i2,t} \quad (A4)$$

and everything is defined analogously to equation (A3) but at the industry level. So, $share_{i3}$, is the average share of public expenditure on defense R&D in the firm's three-digit industry ("i3") as a share of the two-digit industry ("i2"). And $\overline{def}_{i2,t} = def_{i2,(t-1)} + \Delta \overline{def}_{i2,t}$ is the two-digit defense R&D subsidies leaving out the three digit industry subsidies ($\overline{def}_{i2,t} = def_{i2,t} - def_{i3,t}$).

Other data details of the implementation of defense IV

The instrument we use is:

$$Z_{it} = \ln(1 + DR_{it}^{IV})$$

We take logarithms as the R&D series is very skewed and we add a dollar due to zeros (there are no true zeros in the OECD data, so this only affects the French data). We check that the arbitrary adding one did not matter in various ways, such as considering the non-logged version, etc.

Looking at the weight used in (A4) compared to (A2) (or (A3)), note that we could use the US share of industry instead of the French share. We did consider this, but in order to obtain a concordance between France and the US industry definitions we had to aggregate the industries even further, losing much of the industry granularity of the French data.

Note that a practical advantage of our construction of the IV in equation (A2) is that we do not have industry and time specific variation of public defense R&D subsidies for all out countries in the

⁴⁰ One could also have the same concern about *DefGBAORD* in equation (A2) in the OECD data. It would be desirable to remove industry-country specific defense R&D subsidies from the construction of the IV. But, as noted below, we do not observe this for every country-industry-year cell in the OECD. However, nothing critical hinges on this. We can show in the French data, that failing to exclude the own industry defense subsidies only increases the precision of the results to a minor extent.

OECD sample. This was a key motivation for looking at the French data where we can observe defense subsidies at every level from the firm upwards. However, it is also possible to calculate industry specific R&D shares for the US and the UK.

In the US, we can observe data on individual defense contracts that are made publicly available because of procurement transparency rules. Draca (2013) analyses the millions of defense contracts from the early 1960s onwards and generously shared his data with us. His dataset combines historical military procurement data from the National Archives and Records Administration (NARA) with company accounts information from COMPUSTAT. NARA procurement data contains all prime military contracts awarded by the Department of Defense (DoD) between 1966 and 2003. Each year comprises around 250,000 different contracts awarded for the procurement of goods and services, with a minimum reporting threshold of 25,000 USD between 1984 and 2003 (the part of the dataset that we use). The data includes a four-digit product code known as the Federal Supply Code (FSC): we use only expenses with product code “R&D” as our measure of defense related R&D spending. Draca matched the name of the awardee with the COMPUSTAT database. The COMPUSTAT data include a four-digit SIC industry classification (SIC4) for each company. We aggregated the defense R&D expenses by industry and match the SIC4 industry codes to the (more aggregate) industry classification used in the OECD datasets in order to get a distribution of R&D defense spending across industries in the US. The US defense contracts data ends in 2003, so we assume the US weights remain the same for the last few years of the data. The timing of the data is such that it corresponds to the mid-year fiscal year (September) so that the data is partially lagged.

In the UK, we can obtain data from the Office for National Statistics “Expenditure on civil and defense R&D performed in UK businesses by broad product groups,” ONS/GSS reference number: rdbd5, downloaded from <http://statistics.gov.uk/statbase/Product.asp?vlnk=2714&More=Y> in December 2010; now available in the National Archives under <http://webarchive.nationalarchives.gov.uk/20100406130654/http://statistics.gov.uk/statbase/Product.asp?vlnk=2714&More=Y>). There is no UK breakdown prior to 1993 so we assume the weights for the UK were the same in pre-1993 as they were in 1993.

Our baseline specification in the OECD uses US weights for all countries except for the UK (where we use UK weights), but as noted in the main text nothing hinges on this. Our results are robust to dropping the UK and US as well as many other variants of the instrument such as dropping the weights (so setting $w_{it} = 1$), using time invariant weights, using the UK weights for other EU countries instead of the US weights, and so on. This is essentially because the main source of variation comes from the country by year variation.

APPENDIX B: Simulation of a US R&D shock on R&D and productivity, including spillovers

Figure A3 describes the linkages between an R&D shock in the US (calibrated to the 9/11 events) and the way this reverberates throughout the domestic US economy in terms of changes to public and private sector R&D as well as the international economy. We estimated that, on net, such a shock increases US TFP growth by 1.9% and the following calculations explain how this was calculated.

9/11 shock in the US. Between 2001 and 2004 US defense R&D spending (*DefGBAORD*) increased by 52%, from \$46b to \$70b. This represents an increase of US defense R&D spending as a proportion of GDP from 0.45% to 0.60%: by 0.15 percentage points.

Effect on US public R&D. Using the elasticity of public R&D with respect to defense R&D estimated in Table A4, column (3), the first stage of Table 3, column (3), the 9/11 shock generates a 8.7% increase in US public R&D.

Effect on US private R&D (“crowding in”). The increase in public R&D leads to an increase of US private R&D by 3.8%, using the elasticity of public R&D on private R&D as estimated in Table 3, column (3), panel B.

Effect on US total R&D. From Table 1 we can see that in the United States the share of public R&D in total R&D is on average 15.5%. The growth of US public R&D and US private R&D therefore translates into a growth of US total R&D of 4.6%. Since total business conducted R&D (BERD) as a proportion of GDP was 1.95% in the United States in 1997-2001 (Table A2), this increase is equivalent to an increase of BERD/GDP by 0.09 percentage points.

Effect on US TFP growth. Using Table 7, column (3), the increase of R&D/GDP translates into an increase of the TFP growth rate by 0.006 percentage points. Since the US TFP growth rate in 2000 was 0.3%, this represents an increase of around 1.9%.

As a cross check, a simpler method is to take the number can be directly from the reduced form estimates of Table 7, column (4) which has the effect of defense R&D on TFP growth. This produces an estimate of 2.3%.

Hence, both methods suggest that the increased in military R&D on the scale of the 9/11 shock increased TFP growth by about 2%.

So far we have considered only the effect of the 9/11 defense expenditure shock on the US. However, our regressions show that this shock also influences foreign countries. On one hand, there is a displacement effects, lowering private R&D in foreign countries (Table A11).⁴¹ On the other hand, there are direct spillover effects, increasing TFP growth in foreign countries (Table A10). The spillover effects dominate in all our estimations, leading to a net positive increase of TFP growth in the foreign country.

Effect on foreign private R&D (displacement effect). The increase in US total R&D reduces foreign private R&D by the elasticity on the international spillover pool (Table A11, column (3); using geography as distance measure) multiplied by the distance weight of the US to the foreign country. For example, the average geographic distance weight of the US to other countries is 0.02; which translates into reduced private R&D on average of around -0.03% ($=4.6\% * (-0.325) * 0.02$).

Effect on foreign total R&D. Assuming that public R&D is unchanged in the foreign country, this reduction in private R&D is multiplied by the share of private R&D in total R&D (= 1 minus the share of public R&D in total R&D as given by Table 1; for the average country, the share of public R&D in total R&D is 9.5%) for the foreign country to yield the effect on foreign total R&D, which yields a

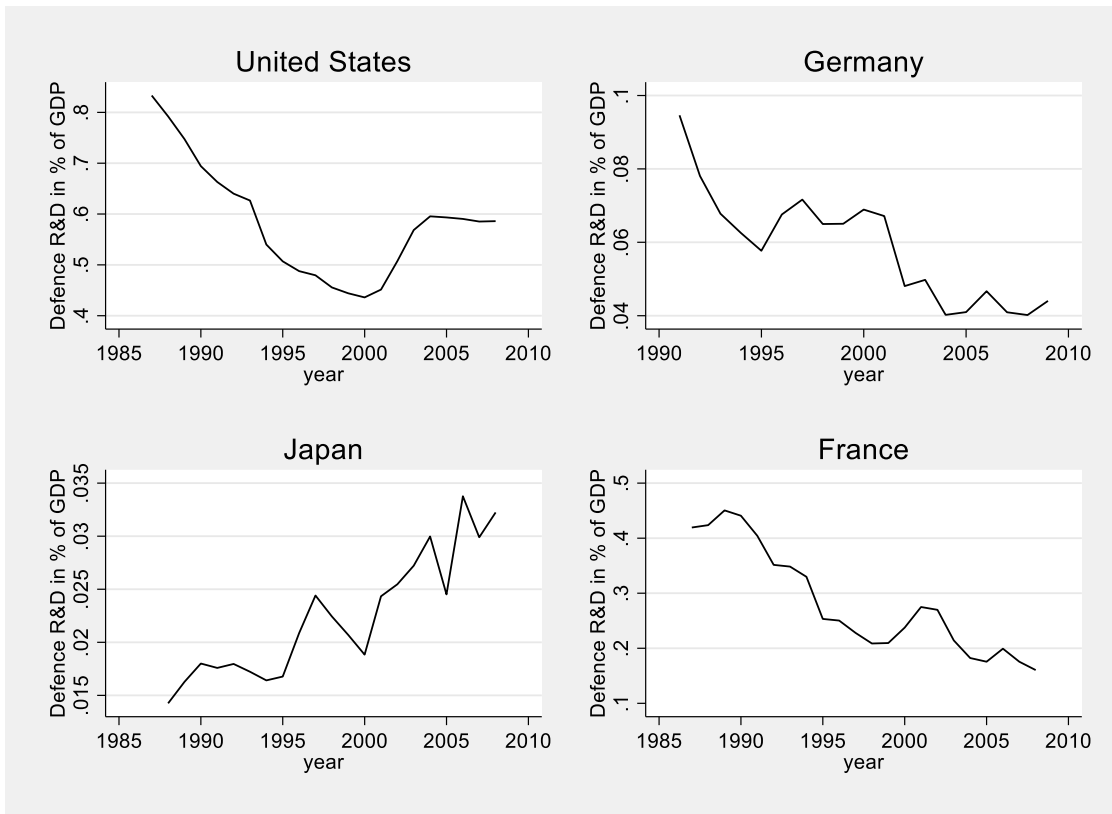
⁴¹ Table A11 is a simplified version of Table 6, which combines the international public and business R&D into one international total R&D spillover pool. This simplifies the calculations in this section.

decrease in total R&D by 0.027%. Table A2 helps again to translate this percentage increase of foreign total R&D into a percentage point increase of foreign total R&D/GDP; for the average country we get a reduction by 0.0003 percentage points in terms of total R&D/VA.

Effect on foreign TFP growth. As discussed above, there is both a displacement effect as well as a positive spillover effect. The displacement effect reduces foreign total R&D as shown above. Using Table A10, reduced foreign R&D/value added leads to reduced foreign TFP growth (given by the estimated coefficient on R&D/value added in column (3) of panel A). For example, in the case of the average country, the TFP growth rate falls by 0.00001 percentage points. However, the increase of US total R&D/VA by 0.09 percentage points (as shown above) has positive spillover effects, as given by the coefficient on the international spillover pool estimated in Table A10, panel A, column (3), again multiplied by the distance weight of the US with respect to the foreign country. In the case of the average country, this leads to an increase in the TFP growth rate by 0.0004 percentage points. Since the positive spillover effect dominates the displacement effect, this is also roughly the same as the net increase in the TFP growth rate; and equivalent to the TFP growth rate increasing by 0.045%.

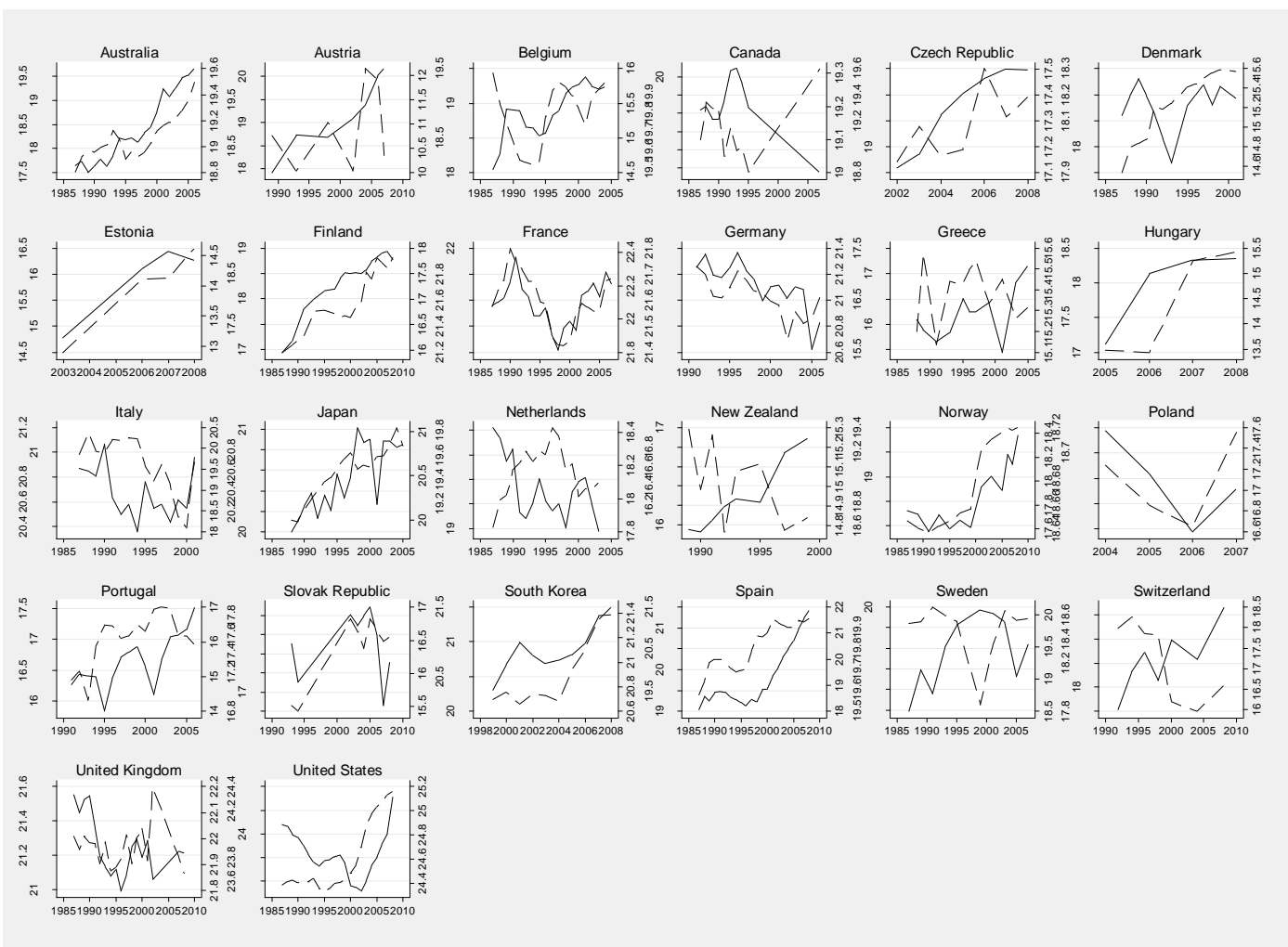
So in conclusion, a 9/11 type shock raises TFP growth by 1.9% in the US and an average of 0.045% in the rest of the OECD. This shock, which raised US defense R&D spending by 52%, would result in an additional \$37 billion (nominal, PPP adjusted) if it happened again in 2013. According to our estimated increase in the US TFP growth rate, this then leads to an additional \$1bn GDP per annum in the US, and an additional \$0.3bn GDP per annum in the rest of the world.

FIGURE 1: DEFENSE R&D AS PERCENT OF GDP IN THE US, GERMANY, JAPAN AND FRANCE



Notes: This figure shows the defense related, government funded total R&D as a share of GDP. The defense R&D in this figure refers to “all public R&D” which includes all government budget appropriations or outlays of total R&D, i.e. not just the government funded part of business conducted R&D, but also the government funded part of R&D conducted outside of enterprises

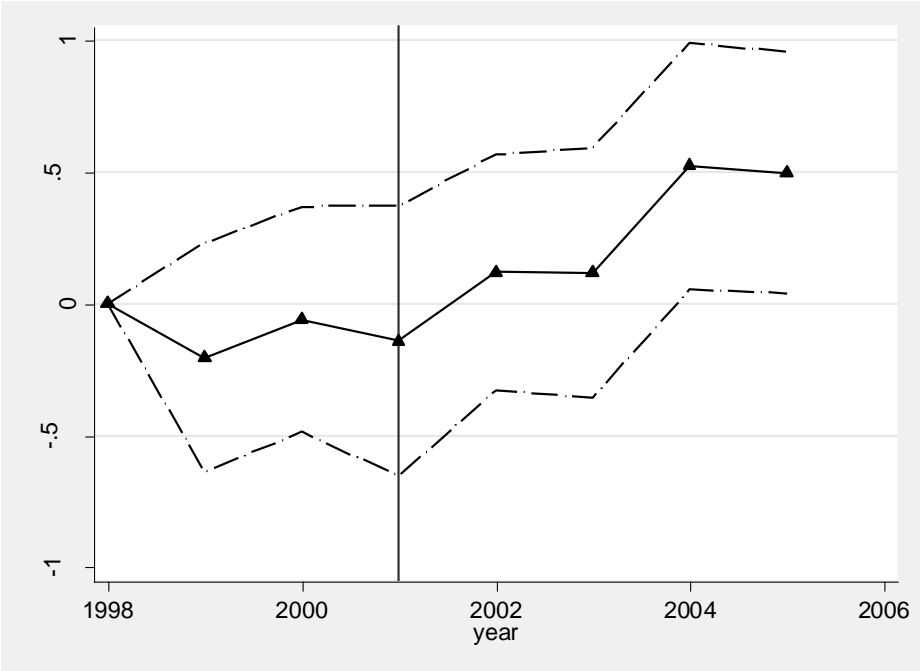
FIGURE 2: PUBLIC R&D AND TOTAL DEFENSE R&D – BY COUNTRY



——— $\ln(\text{government funded and business conducted R\&D})$
 - - - $\ln(\text{defense related total government funded R\&D})$

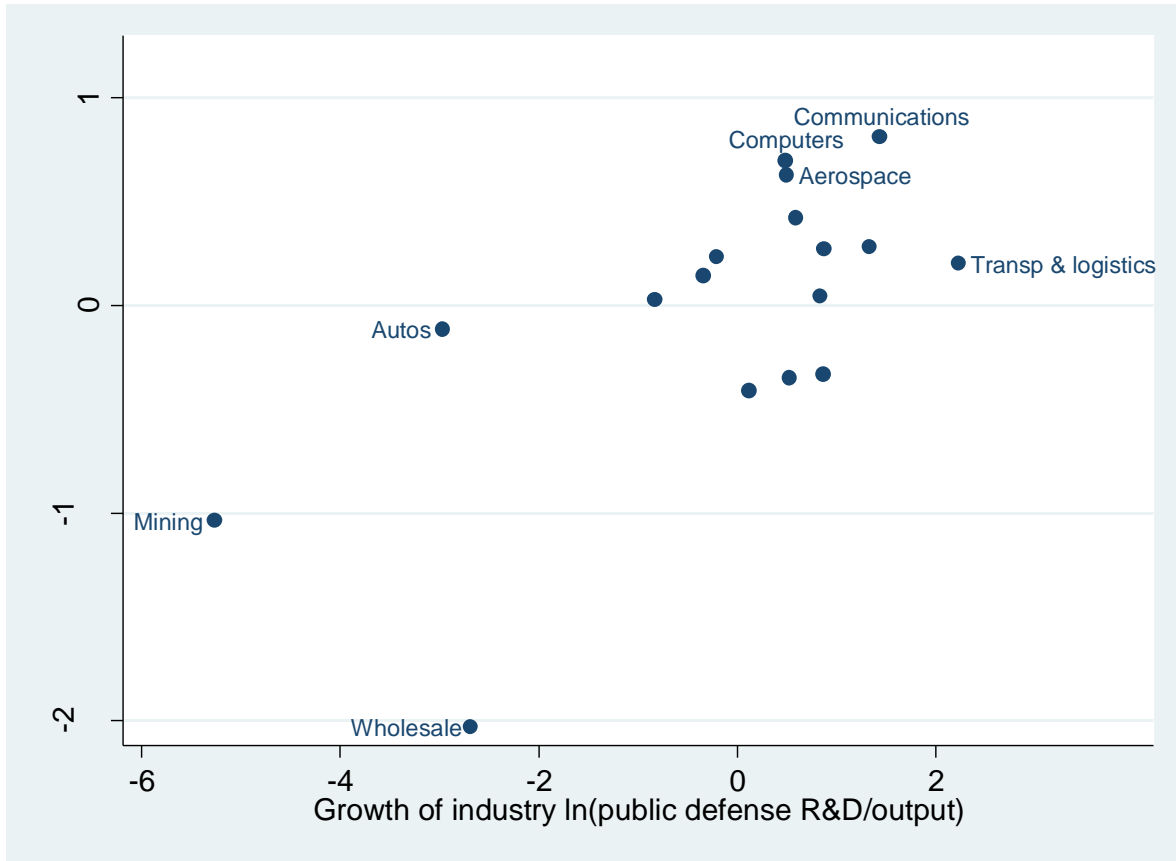
Notes: Solid line is the log of the government funded and business conducted BERD which is scaled on the left axis. The dashed line is the log of government funded total defense R&D (which includes all government budget appropriations or outlays of total R&D, i.e. not just the government funded part of business conducted R&D, but also the government funded part of R&D conducted outside of enterprises), which is scaled on the right axis.

FIGURE 3: PRIVATE R&D INTENSITY IN DEFENSE SENSITIVE SECTORS RELATIVE TO OTHER SECTORS BEFORE AND AFTER 9/11 2001 – US DATA



Notes: These are based on difference in difference regressions 1998-2005. “Defense sensitive sectors” are aerospace and other transport; information technology; and Communication Technologies (Radio, TV and Communications Equipment). We run OLS regressions with industry dummies and time dummies and an interaction between the defense sensitive sectors and different year dummies pre and post 2001. Solid line is the OLS coefficient (base period is 1998) and 95% confidence intervals based on robust standard errors.

FIGURE 4: GROWTH IN PRIVATE R&D INTENSITY AND PREDICTED DEFENSE R&D INTENSITY BEFORE AND AFTER 9/11 2001



Notes: The horizontal axis is the change in industry-level ln(predicted defense R&D/output) pre and post 9/11; the vertical axis is the same for ln(privately funded R&D/output). Pre-policy is 1999 and post-policy is 2005. The correlation is 0.66 and significant at the 1% level.

TABLE 1: PUBLIC R&D AS A PROPORTION OF TOTAL BUSINESS CONDUCTED R&D (BERD) - BY COUNTRY

| Country | Mean | All years Standard deviation | Maximum- Minimum | 1987-1991 | 1992-1996 | 1997-2001 | 2002-2006 | 2007-2009 |
|----------------|-------|------------------------------------|---------------------|-----------|-----------|-----------|-----------|-----------|
| | | | | Mean | Mean | Mean | Mean | Mean |
| Australia | 3.32 | 0.78 | 2.76 | 3.29 | 2.52 | 3.43 | 4.03 | |
| Austria | 7.56 | 2.22 | 4.75 | 5.63 | 9.81 | 5.53 | 7.22 | 10.28 |
| Belgium | 5.52 | 1.16 | 5.05 | 6.00 | 4.89 | 5.86 | 5.52 | |
| Canada | 8.84 | 2.85 | 9.38 | 10.42 | 8.54 | | | 2.18 |
| Czech Rep. | 10.81 | 4.10 | 11.92 | | 4.72 | 11.43 | 13.28 | 13.78 |
| Denmark | 7.21 | 3.07 | 8.76 | 10.54 | 6.09 | 4.16 | | |
| Estonia | 8.86 | 4.12 | 13.76 | | | 10.06 | 7.67 | 8.23 |
| Finland | 3.92 | 0.97 | 3.58 | 3.91 | 5.85 | 3.92 | 3.53 | 3.00 |
| France | 13.64 | 4.60 | 13.9 | 20.88 | 14.09 | 9.53 | 10.86 | 9.78 |
| Germany | 7.96 | 2.24 | 6.25 | 10.06 | 10.39 | 7.65 | 5.42 | |
| Greece | 5.91 | 2.60 | 10.03 | 8.44 | 6.15 | 3.72 | 5.02 | |
| Hungary | 9.54 | 4.79 | 19.01 | 8.17 | 14.70 | 8.37 | 6.01 | 9.11 |
| Italy | 14.36 | 3.25 | 10.42 | 17.83 | 13.12 | 12.57 | 12.16 | |
| Japan | 1.39 | 0.28 | 1.24 | 1.40 | 1.28 | 1.53 | 1.32 | |
| Netherlands | 7.44 | 3.34 | 11.62 | 11.60 | 7.13 | 5.05 | 3.84 | |
| New Zealand | 7.85 | 1.35 | 3.82 | 6.62 | 7.14 | 9.08 | 9.98 | |
| Norway | 12.52 | 3.98 | 11.43 | 18.52 | 13.93 | 10.32 | 9.96 | 9.24 |
| Poland | 23.11 | 8.77 | 22.25 | | 32.00 | 29.01 | 14.15 | 11.68 |
| Portugal | 5.94 | 2.46 | 7.26 | 5.00 | 7.97 | 6.47 | 4.30 | |
| Slovakia | 17.59 | 6.20 | 19.93 | | 11.90 | 20.91 | 23.30 | 10.11 |
| South Korea | 5.53 | 1.18 | 4.51 | | 3.96 | 6.31 | 5.17 | 6.06 |
| Spain | 11.34 | 2.96 | 11.33 | 12.78 | 9.93 | 8.12 | 12.24 | 17.12 |
| Sweden | 8.25 | 2.80 | 8.09 | 11.38 | 10.13 | 7.10 | 5.20 | 4.68 |
| Switzerland | 1.91 | 0.40 | 0.91 | | 2.07 | 2.28 | 1.50 | 1.65 |
| United Kingdom | 11.54 | 3.96 | 13.42 | 17.10 | 10.68 | 9.67 | 7.07 | 6.69 |
| United States | 15.53 | 7.00 | 23.53 | 26.61 | 17.16 | 11.08 | 9.31 | 12.13 |

Notes: Our database comprises of an unbalanced panel of yearly values between the years 1987 and 2009. “Maximum-minimum” is the difference between the highest and the lowest value within a country across all years.

TABLE 2: DEFENSE SHARE OF TOTAL PUBLIC R&D

| Country | All years | All years | 1987-1991 | 1992-1996 | 1997-2001 | 2002-2006 | 2007-2009 |
|----------------|-----------|--------------------|-----------|-----------|-----------|-----------|-----------|
| | Mean | Standard deviation | Mean | Mean | Mean | Mean | Mean |
| Australia | 8.09 | 1.87 | 11.18 | 8.40 | 6.70 | 6.86 | 7.34 |
| Austria | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.01 | 0.00 |
| Belgium | 0.38 | 0.20 | 0.57 | 0.31 | 0.40 | 0.31 | 0.24 |
| Canada | 5.08 | 1.32 | 6.86 | 4.88 | 5.08 | 3.86 | 3.24 |
| Czech Rep. | 2.70 | 0.44 | | | | 3.02 | 2.33 |
| Denmark | 0.60 | 0.21 | 0.44 | 0.61 | 0.53 | 0.88 | 0.55 |
| Estonia | 0.96 | 0.55 | | | | 0.92 | 1.04 |
| Finland | 1.97 | 0.58 | 1.54 | 1.97 | 1.43 | 2.56 | 2.32 |
| France | 28.98 | 6.19 | 37.26 | 32.36 | 23.10 | 23.37 | 28.56 |
| Germany | 8.47 | 2.58 | 12.46 | 9.22 | 8.36 | 6.01 | 5.92 |
| Greece | 1.21 | 0.60 | 1.91 | 1.54 | 0.92 | 0.58 | 0.55 |
| Hungary | 0.38 | 0.25 | | | | 0.10 | 0.56 |
| Italy | 4.47 | 2.83 | 7.65 | 6.45 | 2.62 | 3.29 | 2.16 |
| Japan | 5.03 | 0.74 | 5.27 | 6.01 | 4.74 | 4.57 | 4.44 |
| Netherlands | 2.43 | 0.61 | 2.81 | 3.20 | 2.33 | 1.90 | 1.86 |
| New Zealand | 0.91 | 0.41 | 1.50 | 1.07 | 0.73 | 0.95 | 0.00 |
| Norway | 6.15 | 1.30 | 7.77 | 5.64 | 5.77 | 6.55 | 4.82 |
| Poland | 1.91 | 0.85 | | | | 1.42 | 2.64 |
| Portugal | 0.93 | 0.81 | 0.14 | 1.45 | 1.52 | 1.18 | 0.27 |
| Slovakia | 5.65 | 2.06 | | 3.58 | 6.65 | 7.40 | 4.96 |
| South Korea | 16.59 | 2.69 | | | 19.59 | 14.52 | 17.22 |
| Spain | 17.63 | 7.83 | 15.19 | 11.79 | 27.63 | 20.41 | 10.11 |
| Sweden | 18.08 | 6.28 | 25.11 | 21.69 | 10.27 | 18.65 | 12.41 |
| Switzerland | 4.10 | 3.92 | 10.17 | 4.84 | 2.29 | 0.50 | 0.59 |
| United Kingdom | 35.00 | 8.16 | 44.27 | 39.08 | 36.10 | 28.97 | 20.93 |
| United States | 57.20 | 4.98 | 64.84 | 56.34 | 52.94 | 55.50 | 55.83 |

Note: Our database comprises of an unbalanced panel of yearly values between the years 1987 and 2009. The defense share of public R&D in this table refers to “all public R&D” which includes all government budget appropriations or outlays of total R&D, i.e. not just the government funded part of business conducted R&D, but also the government funded part of R&D conducted outside of enterprises; while public defense R&D is the defense related part of “all public R&D”.

TABLE 3: EFFECT OF PUBLICLY FUNDED R&D ON PRIVATELY FUNDED R&D – OECD DATA

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| Dependent variable: ln(Privately funded business R&D) | | | | | | |
| Panel A: OLS | | | | | | |
| ln(Public R&D) _{t-1} | 0.152** (0.043) | 0.168** (0.041) | 0.164** (0.038) | 0.162** (0.037) | 0.160** (0.035) | 0.160** (0.035) |
| R&D tax credit _{t-1} | | | | 0.955** (0.353) | 0.910** (0.345) | 0.907** (0.343) |
| ln(Non-bus. public R&D) _{t-1} | | | | | 0.072 (0.202) | 0.073 (0.203) |
| (Corp. tax revenue/GDP) _{t-1} | | | | | | 0.120 (3.183) |
| ln(output) _{t-1} | | | 0.859** (0.180) | 0.862** (0.181) | 0.876** (0.188) | 0.877** (0.185) |
| ln(GDP) _{t-1} | | | 0.469 (0.338) | 0.302 (0.347) | 0.252 (0.382) | 0.247 (0.427) |
| Panel B: IV | | | | | | |
| ln(Public R&D) _{t-1} | 0.813** (0.291) | 0.511** (0.207) | 0.434** (0.179) | 0.389** (0.177) | 0.454** (0.224) | 0.457** (0.226) |
| R&D tax credit _{t-1} | | | | 0.818** (0.384) | 0.906** (0.412) | 0.873** (0.408) |
| ln(Non-bus. public R&D) _{t-1} | | | | | -0.197 (0.235) | -0.193 (0.234) |
| (Corp. tax revenue/GDP) _{t-1} | | | | | | 1.144 (3.790) |
| ln(output) _{t-1} | | | 0.650** (0.232) | 0.687** (0.235) | 0.603** (0.302) | 0.609** (0.299) |
| ln(GDP) _{t-1} | | | 0.692* (0.408) | 0.513 (0.425) | 0.706 (0.574) | 0.659 (0.595) |
| Observations | 4,951 | 4,181 | 4,181 | 4,181 | 4,181 | 4,181 |
| First stage F-Statistic | 10.89 | 13.05 | 15.50 | 14.94 | 8.413 | 8.232 |
| Anderson-Rubin | 7.556 | 4.450 | 4.308 | 3.630 | 3.603 | 3.602 |
| Wald F-test p-value | 0.00631 | 0.0358 | 0.0388 | 0.0578 | 0.0587 | 0.0587 |

Notes: Two-way clustered standard errors at the industry-country and country-year level. The dependent variable is private R&D, i.e., R&D conducted in the business sector (BERD) that is also financed by the private sector (i.e., excludes government financed R&D). “Public R&D” is government-financed R&D performed by private firms. “R&D tax credit” is an index from Thomson (2012). “Non-business public R&D” is government-financed R&D performed not by the private sector, e.g. universities or other institutions. “Corporate tax revenue/GDP” is tax revenue from taxes on income, profits, and capital gains of corporates divided by GDP (from OECD). All columns include a full set of country*industry fixed effects and industry*year fixed effects. Models in Panel B use government funded R&D in defense as an instrument for government-financed R&D. The Anderson-Rubin Wald F-test tests the null hypothesis of weak instruments. All columns include weights given by the employment share of an industry within a country in the first year it appears in the sample. We use predicted defense R&D to instrument for public R&D. ** significant at 5% level, * significant at 10% level.

TABLE 4: EFFECT OF PUBLICLY-FUNDED R&D ON PRIVATELY-FUNDED R&D – FRENCH DATA

| | (1) OLS | (2) IV | (3) OLS | (4) IV |
|---|---------------------|--------------------|---------------------|--------------------|
| Dependent variable: ln(Privately funded business R&D) | | | | |
| Panel A. 3-digit industry | | | | |
| ln(public R&D+1) _{t-1} | 0.069*** (0.011) | 0.346** (0.092) | | |
| ln(defense R&D+1) _{t-1} | | | 0.047*** (0.009) | 0.150** (0.041) |
| Observations | 4,444 | 4,444 | 4,444 | 4,444 |
| Number of ind | 169 | 169 | 169 | 169 |
| First stage F | | 11.56 | | 23.23 |
| p(firm β =ind β) | 0.000 | 0.014 | 0.000 | 0.000 |
| Panel B. Firm-level | | | | |
| ln(public R&D+1) _{t-1} | 0.011*** (0.002) | 0.119* (0.069) | | |
| ln(defense R&D+1) _{t-1} | | | 0.006 (0.005) | 0.374* (0.215) |
| Observations | 80,692 | 80,692 | 81,201 | 81,201 |
| Number of units (industries or firms) | 12,429 | 12,429 | 12,586 | 12,586 |
| First stage F | | 13.76 | | 12.12 |
| p(firm β = industry β) | 0.000 | 0.001 | 0.000 | 0.298 |

Notes: The dependent variable is private R&D, i.e., R&D conducted in the business sector (BERD) that is also financed by the private sector (i.e., excludes government financed R&D). “Public R&D” is government-financed R&D performed by private firms. All columns include year fixed effects; Panel A (B) includes 3-digit industry (firm) fixed effects. All variables in Panel B are winsorized at the 0.5% tails. “First stage F” is the Kleibergen-Paap rk Wald F-statistic. We use defense R&D subsidies to instrument for public R&D. Standard errors in Panel A (B) are clustered at the 2-digit (3-digit) industry level. ** significant at 5% level, * significant at 10% level.

TABLE 5: EFFECT OF PUBLICLY FUNDED R&D ON EMPLOYMENT AND WAGES – IV ESTIMATES

| | (1) ln(BERD excl. gov funded) | (2) ln(BERD excl. gov funded) | (3) ln(R&D personnel) | (4) ln(employment excluding R&D) | (5) ln(R&D labor cost/R&D personnel) |
|---|-------------------------------------|-------------------------------------|-----------------------------|--|--|
| Panel A. OECD country-industries | | | | | |
| ln(public R&D) _{t-1} | 0.813** (0.291) | 0.591** (0.294) | 0.563* (0.329) | 0.127** (0.048) | 0.123 (0.121) |
| Observations | 4,951 | 3,690 | 3,690 | 3,690 | 3,690 |
| First stage F | 10.89 | 12.19 | 12.19 | 12.19 | 12.19 |
| Panel B. French 3-digit industries | | | | | |
| ln(public R&D) _{t-1} | 0.346*** (0.092) | 0.403*** (0.090) | 0.394*** (0.069) | 0.324*** (0.086) | 0.009 (0.013) |
| Observations | 4,444 | 3,982 | 3,982 | 3,982 | 3,982 |
| First stage F | 11.56 | 10.40 | 10.40 | 10.40 | 10.40 |
| Panel C. French firms | | | | | |
| ln(public R&D) _{t-1} | 0.119* (0.069) | 0.114 (0.071) | 0.114** (0.049) | 0.019 (0.057) | 0.033 (0.043) |
| Observations | 80,692 | 80,044 | 80,044 | 80,044 | 80,044 |
| First stage F | 13.76 | 13.51 | 13.51 | 13.51 | 13.51 |

Notes: The dependent variable is private R&D, i.e., R&D conducted in the business sector (BERD) that is also financed by the private sector (i.e., excludes government financed R&D). “Public R&D” is government-financed R&D performed by private firms. Panel A (B) includes firm (3-digit industry) and year fixed effects. Panel C includes industry*country and industry*year fixed effects. We use defense R&D subsidies to instrument for public R&D (details are provided in Section 2). All columns of panel A include weights given by the employment share of an industry within a country in the first year it appears in the sample. Standard errors in Panel A are clustered at the 3-digit industry level; in Panel B they are clustered at the 2-digit industry level; in Panel C they are two-way clustered at the country-industry and country-year level. ** significant at 5% level, * significant at 10% level.

TABLE 6: INTERNATIONAL SPILLOVERS – IV ESTIMATES – OECD DATA

| Method of calculating Spillovers from neighbors | (1) Baseline | (2) Per-capita GDP | (3) Geo- graphy | (4) Skill Intensity | (5) Tech- nology | (6) FDI Flows | (7) R&D Intensity |
|--|--------------------|-----------------------|-----------------------|---------------------------|------------------------|---------------------|-------------------------|
| Dependent variable: ln(Privately funded business R&D) | | | | | | | |
| Panel A. Effect of Public R&D | | | | | | | |
| ln(public R&D) _{t-1} | 0.813** (0.291) | 0.893** (0.339) | 0.704** (0.201) | 0.735** (0.238) | 0.735** (0.234) | 0.712** (0.214) | 0.921** (0.367) |
| ln(int. public R&D) _{t-1} | | 0.158* (0.093) | 0.504** (0.174) | 0.150 (0.103) | 1.014** (0.281) | 0.538** (0.195) | 0.263* (0.152) |
| Observations | 4,951 | 4,948 | 4,948 | 4,401 | 4,948 | 4,585 | 4,948 |
| First stage F | 10.89 | 8.768 | 16.01 | 15.68 | 14.57 | 14.06 | 7.925 |
| Panel B. Effect of Business R&D | | | | | | | |
| ln(public R&D) _{t-1} | 0.813** (0.291) | 0.789** (0.289) | 0.810** (0.281) | 0.830** (0.300) | 0.763** (0.269) | 0.776** (0.282) | 0.748** (0.276) |
| ln(int. business R&D) _{t-1} | | -0.052 (0.056) | 0.040 (0.157) | -0.045 (0.098) | -0.703** (0.272) | -0.027 (0.173) | -0.302** (0.114) |
| Observations | 4,951 | 4,948 | 4,948 | 4,401 | 4,948 | 4,585 | 4,948 |
| First stage F | 10.89 | 10.73 | 12.51 | 11.84 | 10.36 | 10.81 | 11.01 |
| Panel C. Effects of Public and Business R&D | | | | | | | |
| ln(public R&D) _{t-1} | 0.813** (0.291) | 0.856** (0.318) | 0.693** (0.211) | 0.799** (0.260) | 0.652** (0.212) | 0.657** (0.197) | 0.876** (0.349) |
| ln(int. public R&D) _{t-1} | | 0.348** (0.152) | 0.686** (0.196) | 0.247* (0.149) | 1.111** (0.295) | 0.690** (0.194) | 0.474** (0.188) |
| ln(int. business R&D) _{t-1} | | -0.287** (0.121) | -0.407** (0.164) | -0.216 (0.154) | -1.066** (0.259) | -0.520** (0.190) | -0.616** (0.186) |
| Observations | 4,951 | 4,948 | 4,948 | 4,401 | 4,948 | 4,585 | 4,948 |
| First stage F | 10.89 | 8.915 | 15.53 | 14.29 | 14.80 | 14.25 | 8.186 |

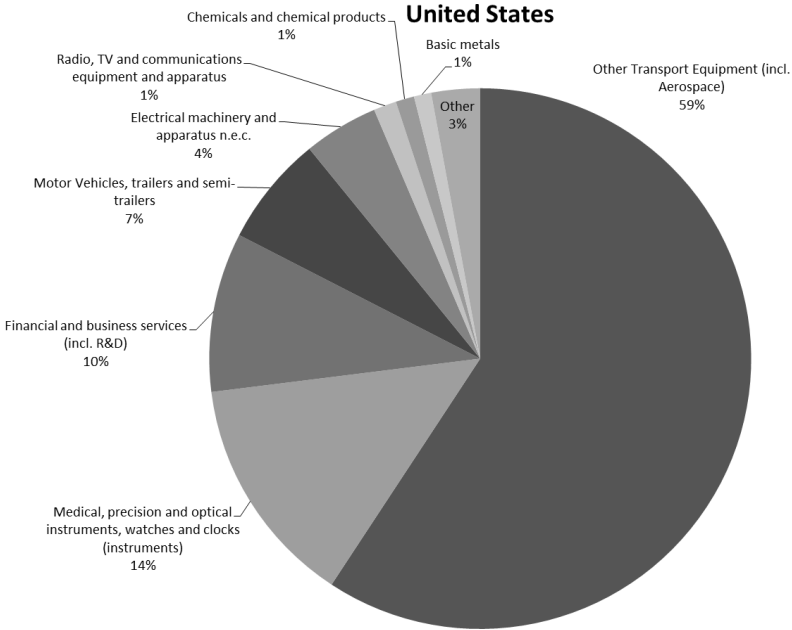
Notes: Each column in each panel presents a separate regression. Two-way clustered standard errors at the industry-country and country-year level. All columns include a full set of country*industry fixed effects and industry*year fixed effects. All columns include weights given by the employment share of an industry within a country in the first year it appears in the sample. The dependent variable is private R&D. Public R&D is government-financed R&D performed by private firms. “Int. public R&D” is the weighted average of other countries’ public R&D in the same industry and year, where each column uses different weights. Weights ω_{kl} are a “distance” measure between country k and country l measured by: (2) GDP per capita. $\omega_{kl} = 1/|GDPcap_k - GDPcap_l|$; (3) Geographic distance. $\omega_{kl} = 1/dist_{kl}$ where geographic distance is measured in kilometers between capital cities; (4) Skill intensity. $\omega_{kl} = 1/|tert_k - tert_l|$ with $tert_t$ being the share of population with tertiary education in country l ; (5) Patent similarity. $\omega_{kl} = \sqrt{\sum_i (pat_{ik} - pat_{il})^2}$, pat_{ik} and pat_{il} are patent share of a specific technology class I (out of 15 patent technology classes) in country k or l (of total patents in that country); (6) Inward FDI flows. $\omega_{kl} = \frac{FDI_{kl}}{\sum_{j=1}^J FDI_{kj}}$ where FDI_{kl} is the amount of FDI that country k receives from country l ; (7) R&D intensity. $\omega_{kl} = 1/\left|\frac{Total\ R\&D_k}{GDP_k} - \frac{Total\ R\&D_l}{GDP_l}\right|$. ** significant at 5% level, * significant at 10% level.

TABLE 7: EFFECT OF R&D GROWTH ON TFP GROWTH

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|--|--------------------|--------------------|--------------------|--------------------|----------------------------------|--------------------|--------------------|--------------------|
| | OLS | OLS | IV | RF | OLS | OLS | OLS | OLS |
| | Annual TFP growth | | | | Annual labor productivity growth | | | |
| (Total R&D/value added) _{t-1} | 0.097** (0.042) | 0.099** (0.043) | 0.064** (0.027) | | | | | |
| (Predicted Defense R&D/value added) _{t-1} | | | | 0.049** (0.017) | | | | |
| (Total R&D/ sales) _{t-1} | | | | | 0.026** (0.004) | 0.026** (0.004) | 0.040** (0.004) | 0.043** (0.004) |
| Observations | 4,779 | 4,779 | 4,779 | 4,779 | 4,404 | 4,404 | 82,743 | 82,743 |
| Dataset | OECD | OECD | OECD | OECD | French Industries | French Industries | French Firms | French Firms |
| First stage F | | | 18.72 | | | | | |
| Country FE | NO | YES | YES | YES | N/A | N/A | N/A | N/A |
| 2-digit Industry FE | N/A | N/A | N/A | N/A | NO | YES | NO | YES |

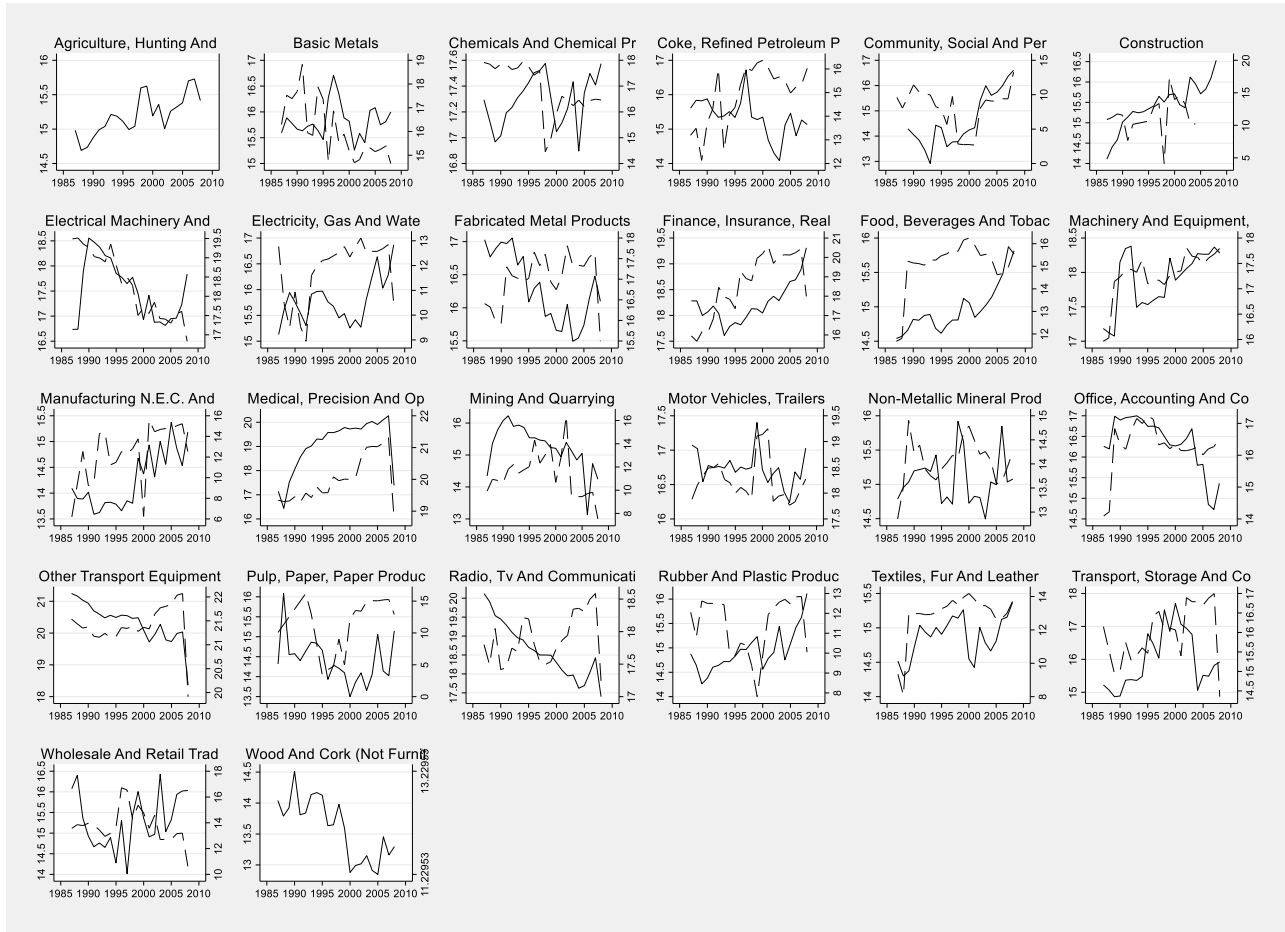
Notes: Total R&D is the sum of private R&D, public R&D and R&D from any other source. Columns (1) through (4) use two-way clustered standard errors at the industry*country and country*year level, and include weights given by the employment share of an industry within a country in the first year it appears in the sample. Columns (5) through (8) use clustered standard errors at the 3-digit industry. All regressions include a full set of year fixed effects. Column (3) presents IV estimates using (defense R&D)/value added as an instrument for R&D/value added. Column (4) presents the reduced form estimate for column (3). ** significant at 5% level, * significant at 10% level.

APPENDIX FIGURE A1: INDUSTRY SHARES OF DEFENSE R&D IN THE US



Notes: This figure shows each industry’s share of defense R&D averaged over 1987-2003. US data are from Department of Defense prime procurement contracts, generously provided by Mirko Draca (see Draca, 2012 for exact construction).

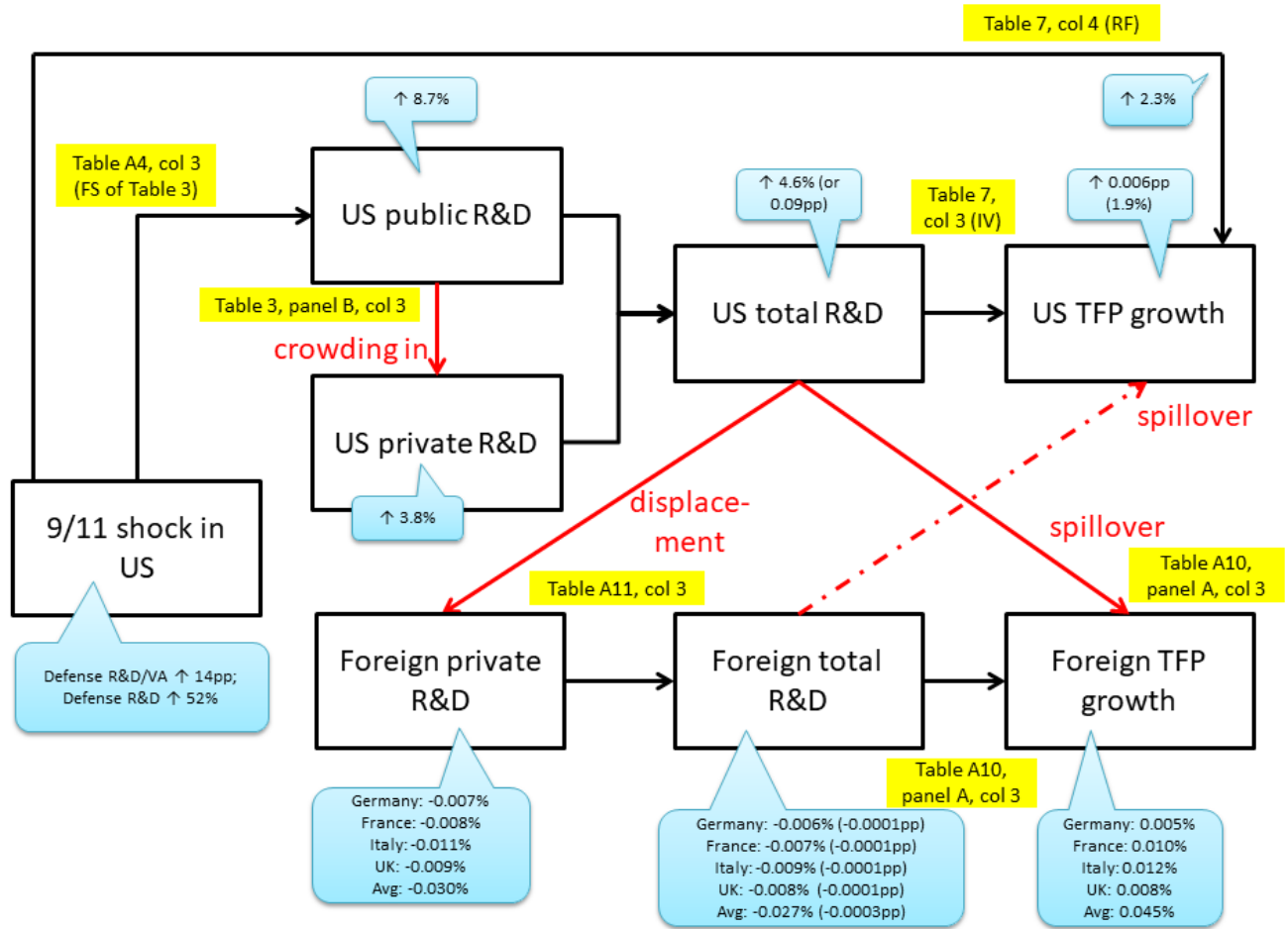
APPENDIX FIGURE A2: PUBLIC R&D AND PREDICTED DEFENSE R&D – BY INDUSTRY



—— ln(government funded and business conducted R&D)
 - - - ln(predicted defense related government funded R&D)

Notes: The solid line is the log of the government funded and business conducted BERD which is scaled on the left axis. The dashed line is the log of predicted defense R&D, our instrumental variable as explained in the text, which is scaled on the right axis. These are averages across all countries in our data set. Agriculture, hunting and forestry has zero predicted defense R&D in all years. Wood and cork has zero predicted defense R&D in all years but one.

APPENDIX FIGURE A3: THE EFFECTS OF A 9/11 SHOCK OF US DEFENSE R&D ON US AND FOREIGN TFP GROWTH AND PUBLIC AND PRIVATE R&D



Note: This describes the links from an exogenous increase in US defense R&D (calibrated to the increase after 9/11) to TFP growth in the US and overseas. Details are in Appendix B.

APPENDIX TABLE A1: VARIABLE DEFINITION AND SUMMARY STATISTICS
PANEL A – OECD DATA

| Variable | Code | Construction | Source | Unit | Mean | Median | SD |
|------------------------------|-------------------|--|------------|--------------------|--------|--------|--------|
| R&D | R | Business enterprise R&D (BERD): Total R&D conducted by businesses from any funding source (government, business, overseas) | MSTI | 2000\$ PPP adj, mn | 898 | 92 | 2,868 |
| Public R&D | S | Government funded part of business conducted R&D (BERD) | MSTI, SIRD | 2000\$ PPP adj, mn | 71 | 4 | 451 |
| IV for defense R&D subsidies | DR ^{IV} | Predicted public subsidies for Defense R&D (see Appendix A for details) | MSTI | 2000\$ PPP adj, mn | 159 | 0.2 | 1,645 |
| # Scientists | | Number of R&D scientists | MSTI | FTE | 5,699 | 564 | 16,943 |
| R&D personnel | L ^R | Total R&D workers | MSTI | FTE | 5,143 | 884 | 14,226 |
| R&D wage bill | (WL) ^R | Labor costs of all R&D personnel | MSTI | 2000\$ PPP adj, mn | 276 | 41 | 790 |
| R&D wage | W ^R | BERD labor cost/number of R&D personnel | MSTI | 2000\$ PPP adj | 50,650 | 50,886 | 20,669 |
| Output | Y | Total production (gross output) | STAN | volumes, bn | 6,914 | 36 | 27,085 |
| GDP | GDP | Total production (gross output) of a country | STAN | 2000\$ PPP adj, bn | 1,180 | 485 | 2,028 |
| Wage bill | WL | Total cost of all employees | STAN | 2000\$ PPP adj, mn | 12,003 | 1,993 | 61,921 |
| Employment | L | Number of persons engaged (total employment) | STAN | thousand, persons | 499 | 104 | 1,850 |
| Value Added | VA | Value added | STAN | volumes, bn | 2,117 | 11 | 11,780 |
| Capital | K | Gross capital stock; if missing in STAN own estimation based on investment | STAN | volumes, bn | 4,153 | 36 | 47,879 |

PANEL B – FRENCH DATA

| Variable | Code | Construction | Firms | | | Industries | | |
|---------------|-------------------|---|---------|---------|---------|------------|---------|----------|
| | | | Mean | Median | SD | Mean | Median | SD |
| R&D | R | Business enterprise R&D (BERD): | €8.3m | €0.8m | €56.9m | €62.1m | €4.7m | €226.5m |
| Public R&D | S | Government funded part of business conducted R&D (BERD) | €0.84m | €0.0 | €13.4m | €6.3m | €0.0 | €49.0m |
| Defense R&D | DR | Defense R&D | €0.6m | €0.0 | €12m | €4.3m | €0.0 | €42.8m |
| # Scientists | | Number of R&D scientists (FTE) | 29 | 5 | 171 | 218 | 19 | 776 |
| R&D personnel | L ^R | Total R&D workers (FTE) | 62 | 10 | 363 | 463 | 49 | 1526 |
| R&D wage bill | (WL) ^R | Labor costs of all R&D personnel | €3.5m | €0.5m | €20.9m | €26.3m | €2.5m | €90.0m |
| R&D wage | W ^R | BERD labor cost/ number of R&D personnel | €55,313 | €50,656 | €30,070 | €53,370 | €50,341 | €31,086 |
| Output | Y | Sales | €260m | €19m | €6,766m | €1,945m | €282 | €18,753m |
| Employment | L | Number of persons engaged (total employment), FTE | 925 | 129 | 7348 | 6,939 | 1,376 | 21,991 |

Notes: Summary statistics are based on the sample used in Table 2 (OECD data in Panel A) and Table 4 (French data in Panel B). All values in 2000 US dollars (Panel A) and 2000 Euros (Panel B) unless otherwise stated. All data is from the Research Ministry R&D survey.

APPENDIX TABLE A2: BUSINESS CONDUCTED R&D (BERD) AS A PROPORTION OF GDP - BY COUNTRY

| Country | All years | | | 1987-1991 | 1992-1996 | 1997-2001 | 2002-2006 | 2007-2009 |
|----------------|-----------|--------------------|-----------------|-----------|-----------|-----------|-----------|-----------|
| | Mean | Standard deviation | Maximum-Minimum | Mean | Mean | Mean | Mean | Mean |
| Australia | 0.78 | 0.24 | 0.74 | 0.48 | 0.73 | 0.88 | 1.04 | |
| Austria | 1.27 | 0.39 | 0.96 | 0.81 | 0.73 | 1.11 | 1.57 | 1.50 |
| Belgium | 1.26 | 0.27 | 0.95 | 1.13 | 1.04 | 1.52 | 1.31 | |
| Canada | 0.85 | 0.16 | 0.43 | 0.74 | 0.99 | | | 0.92 |
| Czech Republic | 1.64 | 0.35 | 1.50 | | 1.84 | 1.81 | 1.54 | 1.27 |
| Denmark | 0.89 | 0.36 | 1.06 | 0.63 | 0.72 | 1.19 | 1.65 | |
| Estonia | 0.51 | 0.18 | 0.53 | | | 0.36 | 0.58 | 0.69 |
| Finland | 1.58 | 0.64 | 1.87 | 0.77 | 1.15 | 2.10 | 2.14 | 2.00 |
| France | 1.25 | 0.18 | 0.62 | 1.17 | 1.16 | 1.41 | 1.28 | 1.07 |
| Germany | 1.52 | 0.29 | 0.95 | 1.50 | 1.19 | 1.66 | 1.70 | |
| Greece | 0.19 | 0.06 | 0.20 | 0.12 | 0.14 | 0.23 | 0.23 | |
| Hungary | 0.72 | 0.14 | 0.50 | 1.07 | 0.67 | 0.74 | 0.69 | 0.69 |
| Italy | 0.63 | 0.07 | 0.23 | 0.67 | 0.57 | 0.63 | 0.68 | |
| Japan | 1.53 | 0.34 | 1.14 | 1.42 | 1.17 | 1.61 | 2.03 | |
| Netherlands | 1.02 | 0.16 | 0.50 | 1.09 | 0.85 | 1.16 | 0.98 | |
| New Zealand | 0.38 | 0.13 | 0.41 | 0.30 | 0.33 | 0.46 | 0.55 | |
| Norway | 0.69 | 0.10 | 0.36 | 0.66 | 0.68 | 0.81 | 0.69 | 0.58 |
| Poland | 0.44 | 0.16 | 0.45 | | 0.58 | 0.56 | 0.30 | 0.25 |
| Portugal | 0.23 | 0.11 | 0.43 | 0.17 | 0.14 | 0.25 | 0.37 | |
| Slovakia | 1.01 | 0.67 | 2.49 | | 1.58 | 1.39 | 0.61 | 0.27 |
| South Korea | 2.71 | 0.47 | 1.71 | | 1.87 | 2.67 | 2.88 | 3.26 |
| Spain | 0.56 | 0.13 | 0.38 | 0.46 | 0.42 | 0.61 | 0.68 | 0.71 |
| Sweden | 2.03 | 0.60 | 2.33 | 1.37 | 1.84 | 2.68 | 2.21 | 1.88 |
| Switzerland | 1.44 | 0.21 | 0.51 | | 1.22 | 1.71 | 1.52 | 1.54 |
| United Kingdom | 1.24 | 0.19 | 0.77 | 1.43 | 1.26 | 1.18 | 1.23 | 0.90 |
| United States | 1.87 | 0.09 | 0.34 | 1.89 | 1.78 | 1.95 | 1.81 | 1.97 |

Note: Our database comprises of an unbalanced panel of yearly values between the years 1987 and 2009.

TABLE A3: BUSINESS CONDUCTED R&D AS A PROPORTION OF VALUE ADDED - BY INDUSTRY

| Industry | All years Mean | All years SD | 1987-1991 Mean | 1992-1996 Mean | 1997-2001 Mean | 2002-2006 Mean | 2007-2009 Mean |
|--|---------------------------|-------------------------|---------------------------|---------------------------|---------------------------|---------------------------|---------------------------|
| Agriculture, hunting and forestry | 0.30 | 0.36 | 0.23 | 0.26 | 0.34 | 0.35 | 0.29 |
| Basic metals | 3.81 | 14.25 | 2.19 | 1.98 | 7.36 | 3.17 | 1.35 |
| Construction | 0.18 | 0.27 | 0.10 | 0.14 | 0.22 | 0.25 | 0.21 |
| Chemicals and chemical products | 10.84 | 6.63 | 8.92 | 9.38 | 11.65 | 12.29 | 13.67 |
| Coke, refined petroleum products and nuclear fuel | 5.60 | 18.74 | 3.73 | 4.23 | 4.35 | 4.73 | 29.73 |
| Community, social & personal services, etc. | 0.04 | 0.08 | 0.04 | 0.03 | 0.05 | 0.04 | 0.06 |
| Electricity, gas and water supply | 0.54 | 0.89 | 0.61 | 0.48 | 0.52 | 0.39 | 1.22 |
| Electrical machinery and apparatus n.e.c | 8.68 | 17.02 | 5.98 | 6.92 | 12.26 | 8.70 | 5.22 |
| Finance, insurance, real estate and business activities | 1.17 | 2.69 | 0.42 | 0.87 | 1.10 | 1.38 | 3.21 |
| Fabricated metal products, except machinery and equipment | 0.99 | 0.79 | 0.89 | 0.94 | 1.01 | 1.01 | 1.40 |
| Food, beverages and tobacco | 1.23 | 1.63 | 0.92 | 0.90 | 1.27 | 1.41 | 2.88 |
| Mining and quarrying | 0.98 | 1.25 | 0.88 | 1.01 | 1.09 | 0.96 | 0.59 |
| Machinery and equipment, n.e.c. | 4.91 | 4.08 | 3.66 | 4.31 | 5.21 | 5.56 | 7.42 |
| Manufacturing n.e.c. and recycling | 1.17 | 1.28 | 1.15 | 0.90 | 1.22 | 1.31 | 1.54 |
| Medical, precision and optical instruments, watches and clocks (instruments) | 11.85 | 11.98 | 8.32 | 10.64 | 12.37 | 13.15 | 17.98 |
| Motor Vehicles, trailers and semi-trailers | 10.16 | 10.01 | 11.15 | 8.99 | 9.85 | 10.15 | 14.55 |
| Non-metallic mineral products | 1.31 | 0.94 | 1.24 | 1.14 | 1.43 | 1.44 | 1.19 |
| Office, accounting and computing machinery | 25.12 | 56.00 | 32.57 | 19.76 | 15.91 | 36.31 | 24.35 |
| Other Transport Equipment | 11.36 | 11.27 | 11.01 | 11.50 | 12.80 | 10.47 | 8.78 |
| Pulp, paper, paper products, printing and publishing | 0.59 | 0.51 | 0.48 | 0.58 | 0.61 | 0.63 | 0.60 |
| Radio, TV, communications equipment & apparatus | 23.67 | 15.96 | 24.49 | 21.85 | 24.80 | 24.34 | 20.20 |
| Rubber and plastic products | 2.65 | 2.26 | 1.69 | 2.69 | 3.00 | 2.74 | 3.52 |
| Textiles, fur and leather | 1.00 | 0.86 | 0.58 | 0.74 | 1.06 | 1.32 | 1.93 |
| Transport, storage and communications | 0.47 | 0.95 | 0.24 | 0.31 | 0.57 | 0.44 | 1.33 |
| Wholesale and retail trade; restaurants and hotels | 0.24 | 0.34 | 0.11 | 0.16 | 0.26 | 0.29 | 0.42 |
| Wood and cork (not furniture) | 0.47 | 0.71 | 0.45 | 0.36 | 0.49 | 0.47 | 0.93 |

Note: Our database comprises of an unbalanced panel of yearly values between the years 1987 and 2009. These are averages across all countries in our dataset. SD is for standard deviation.

APPENDIX TABLE A4: FIRST STAGE OF TABLE 3, PANEL B

| | (1) | (2) | (3) | (4) | (5) | (6) |
|---|--------------------|--------------------|--------------------|---------------------|---------------------|---------------------|
| Dependent variable: ln(Privately funded business R&D) | | | | | | |
| ln(predicted defense R&D) _{t-1} | 0.157** (0.045) | 0.163** (0.043) | 0.167** (0.040) | 0.164** (0.040) | 0.120** (0.039) | 0.119** (0.039) |
| R&D tax credit _{t-1} | | | | 0.350 (0.727) | -0.097 (0.740) | 0.001 (0.712) |
| ln(Non-bus. public R&D) _{t-1} | | | | | 0.797** (0.288) | 0.781** (0.279) |
| (Corp. tax revenue/GDP) _{t-1} | | | | | | -3.323 (5.257) |
| ln(output) _{t-1} | | | 0.780** (0.332) | 0.780** (0.329) | 0.913** (0.332) | 0.890** (0.327) |
| ln(GDP) _{t-1} | | | -0.935* (0.512) | -0.993** (0.495) | -1.507** (0.496) | -1.363** (0.485) |
| Observations | 4,951 | 4,181 | 4,181 | 4,181 | 4,181 | 4,181 |

Notes: Two-way clustered standard errors at the industry*country and country*year level. All columns include a full set of country by industry fixed effects and a full set of industry by year fixed effects. All columns use weights given by the employment share of an industry within a country in the first year it appears in the sample. ** significant at 5% level, * significant at 10% level.

APPENDIX TABLE A5: EFFECT OF PUBLIC R&D ON PRIVATE R&D – ROBUSTNESS CHECKS

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
|------------------------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|--------------------|
| | OLS | IV | OLS | IV | OLS | IV | OLS | IV |
| ln(Public R&D) _{t-1} | 0.169** (0.039) | 0.423** (0.163) | 0.154** (0.034) | 0.516** (0.180) | 0.227** (0.035) | 0.570** (0.146) | 0.193** (0.046) | 0.388* (0.211) |
| ln(output) _{t-1} | 0.841** (0.164) | 0.649** (0.208) | -0.028 (0.237) | 0.053 (0.296) | 0.656** (0.179) | 0.301 (0.234) | 1.501** (0.338) | 1.211** (0.453) |
| Left wing govt. _{t-1} | 0.217** (0.079) | 0.107 (0.108) | | | | | | |
| Right wing govt. _{t-1} | 0.238** (0.087) | 0.121 (0.106) | | | | | | |
| Observations | 4,447 | 4,447 | 3,976 | 3,976 | 5,534 | 5,534 | 2,106 | 2,106 |
| Number of ind*ctry | 399 | 399 | 374 | 374 | 456 | 456 | 265 | 265 |
| IV | | Defense | | Defense | | Country Defense | | Defense |
| 1st Stage F | | 17.08 | | 11.66 | | 9.83 | | 4.33 |
| Anderson- Rubin | | 4.556 | | 5.843 | | 6.573 | | 2.19 |
| Wald F-test | | | | | | | | |
| p-value | | 0.0337 | | 0.0163 | | 0.0108 | | 0.141 |

Notes: All columns include a full set of country*industry fixed effects and industry*year fixed effects unless otherwise noted. Columns (1) and (2) present OLS and IV estimates using dummies to control for chief executive’s party orientation, where “center” is the omitted category. Columns (3) and (4) present OLS and IV estimates controlling for ln(output) and ln(GDP) in times t through $t+3$. Columns (5) and (6) present OLS and IV estimates using only country*year variation in the defense instrument. Columns (7) and (8) present OLS and IV estimates controlling for ln(non-R&D military expenditure) in times t through $t+3$. All columns use weights given by the employment share of an industry within a country in the first year it appears in the sample. ** significant at 5% level, * significant at 10% level.

APPENDIX TABLE A6: EFFECT OF PUBLIC R&D AND PRIVATE R&D ON PLACEBO IVS

| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
|---|----------------------------------|---------------------------------|---------------------------------|----------------------------------|---------------------------------|---------------------------------|----------------------------------|---------------------------------|---------------------------------|----------------------------------|---------------------------------|---------------------------------|
| | ln(public R&D) _{t-1} | ln(BERD excl. gov funded) | ln(BERD excl. gov funded) | ln(public R&D) _{t-1} | ln(BERD excl. gov funded) | ln(BERD excl. gov funded) | ln(public R&D) _{t-1} | ln(BERD excl. gov funded) | ln(BERD excl. gov funded) | ln(public R&D) _{t-1} | ln(BERD excl. gov funded) | ln(BERD excl. gov funded) |
| ln(defense proc. excl. R&D) _{t-1} | 0.006 (0.634) | -0.155 (0.326) | | | | | | | | | | |
| ln(defense proc. excl. R&D and other) _{t-1} | | | | 0.753 (0.514) | 0.044 (0.355) | | | | | | | |
| ln(military wagebill excl. R&D) _{t-1} | | | | | | | 0.079 (0.321) | 0.177 (0.182) | | | | |
| ln(mil. wagebill excl. R&D and other) _{t-1} | | | | | | | | | | 0.064 (0.320) | 0.173 (0.178) | |
| ln(public R&D) _{t-1} | | | -24.080 (2,415.548) | | | 0.058 (0.454) | | | 2.239 (7.703) | | | 2.712 (11.935) |
| ln(output) _{t-1} | 1.565** (0.700) | 1.181** (0.410) | 38.871 (3,778.195) | 1.569** (0.697) | 1.166** (0.413) | 1.075 (0.723) | 1.840** (0.787) | 1.307** (0.500) | -2.812 (13.834) | 1.840** (0.796) | 1.287** (0.505) | -3.703 (21.512) |
| ln(GDP) _{t-1} | 0.180 (0.791) | 0.576 (0.521) | 4.916 (443.599) | -0.111 (0.785) | 0.536 (0.543) | 0.542 (0.506) | -0.244 (0.806) | 0.295 (0.576) | 0.840 (2.282) | -0.306 (0.810) | 0.288 (0.575) | 1.118 (3.890) |
| Observations | 2,950 | 2,950 | 2,950 | 2,929 | 2,929 | 2,929 | 2,405 | 2,405 | 2,405 | 2,338 | 2,338 | 2,338 |
| R-squared | 0.037 | 0.058 | -1,076.825 | 0.043 | 0.059 | 0.102 | 0.040 | 0.064 | -8.130 | 0.039 | 0.063 | -12.222 |
| Number of ind*ctry | 331 | 331 | 331 | 326 | 326 | 326 | 298 | 298 | 298 | 275 | 275 | 275 |
| 1st Stage F | | | 8.59e-05 | | | | 1.787 | | 0.0518 | | | 0.0335 |
| Anderson-Rubin Wald | | | 0.189 | | | 0.0126 | | | 0.811 | | | 0.801 |
| F-test | | | | | | | | | | | | |
| p-value | | | 0.664 | | | 0.911 | | | 0.369 | | | 0.372 |

Notes: All columns include a full set of country*industry fixed effects and industry*year fixed. All columns use weights given by the employment share of an industry within a country in the first year it appears in the sample. “Defense procurement excluding R&D and other” excludes R&D, civil defense and foreign military aid. “Military wagebill excluding R&D and other” also excludes R&D, civil defense and foreign military aid. ** significant at 5% level, * significant at 10% level.

APPENDIX TABLE A7: FIRST STAGE OF TABLE 4 – FRENCH DATA

| | (1) | (2) |
|---|--------------------------------------|---------------------------------------|
| | ln(public R&D + 1) _{t-1} | ln(defense R&D + 1) _{t-1} |
| Panel A. 3-digit industry | | |
| ln(predicted defense R&D) _{t-1} | 0.120*** (0.035) | 0.278*** (0.057) |
| Observations | 4,444 | 4,444 |
| Panel B. Firm-level | | |
| ln(predicted defense R&D) _{t-1} | 0.018*** (0.005) | 0.006*** (0.002) |
| Observations | 80,692 | 81,201 |

Notes: Standard errors in Panel A are clustered at the 2-digit industry level; in Panel B are clustered at the 3-digit industry level. All columns include year fixed effects; Panel A includes 3-digit industry fixed effects. Panel B includes firm fixed effects. Variables in panel B are winsorized at the 0.5% tails. The instrument is defense R&D subsidies. ** significant at 5% level, * significant at 10% level.

APPENDIX TABLE A8: EFFECT OF PUBLICLY-FUNDED R&D ON EMPLOYMENT AND WAGES – OLS ESTIMATES OF TABLE 5

| | (1) ln(BERD excl. gov funded) | (2) ln(R&D personnel) | (3) ln(employe nt excluding R&D) | (4) ln(R&D labor cost/R&D personnel) |
|---|-------------------------------------|-----------------------------|---|---|
| Panel A. OECD country-industries | | | | |
| ln(public R&D) _{t-1} | 0.101** (0.042) | 0.110** (0.049) | 0.013** (0.005) | 0.055* (0.032) |
| Observations | 3,690 | 3,690 | 3,690 | 3,690 |
| Panel B. French 3-digit industries | | | | |
| ln(public R&D) _{t-1} | 0.403*** (0.090) | 0.394*** (0.069) | 0.324*** (0.086) | 0.006*** (0.002) |
| Observations | 3,982 | 3,982 | 3,982 | 3,982 |
| Panel C. French firms | | | | |
| ln(public R&D) _{t-1} | 0.011*** (0.002) | 0.013*** (0.001) | 0.001 (0.001) | 0.004*** (0.001) |
| Observations | 80,044 | 80,044 | 80,044 | 80,044 |

Notes: ** significant at 5% level, * significant at 10% level. Panel A (B) is clustered at the 3-digit (2-digit) industry level; Panel C is two-way clustered at the country-industry and country-year level. The dependent variable is private R&D, i.e., R&D conducted in the business sector (BERD) that is also financed by the private sector (i.e., excludes government financed R&D). “Public R&D” is government-financed R&D performed by private firms. Panel A (B) includes firm (3-digit industry) and year fixed effects. Panel C includes industry*country and industry*year fixed effects. All columns in panel A use weights given by the employment share of an industry within a country in the first year it appears in the sample.

APPENDIX TABLE A9: DOMESTIC AND INTERNATIONAL SPILLOVERS – OLS ESTIMATES

| | (1) Baseline | (2) Per-capita GDP | (3) Geo- graphy | (4) Skill Intensity | (5) Tech- nology | (6) FDI Flows | (7) R&D Intensity |
|--|--------------------|--------------------------|-----------------------|---------------------------|------------------------|---------------------|-------------------------|
| Dependent variable: ln(Privately funded business R&D) | | | | | | | |
| Panel A. Effect of Public R&D | | | | | | | |
| ln(public R&D) _{t-1} | 0.152** (0.043) | 0.152** (0.044) | 0.172** (0.042) | 0.151** (0.045) | 0.185** (0.044) | 0.204** (0.042) | 0.151** (0.044) |
| ln(int. public R&D) _{t-1} | | -0.007 (0.039) | 0.264** (0.124) | 0.041 (0.059) | 0.414** (0.143) | 0.214* (0.112) | -0.017 (0.056) |
| Observations | 4,951 | 4,948 | 4,948 | 4,401 | 4,948 | 4,585 | 4,948 |
| Panel B. Effect of Business R&D | | | | | | | |
| ln(public R&D) _{t-1} | 0.152** (0.043) | 0.151** (0.043) | 0.151** (0.043) | 0.147** (0.046) | 0.142** (0.040) | 0.164** (0.043) | 0.153** (0.041) |
| ln(int. business R&D) _{t-1} | | -0.074* (0.042) | -0.082 (0.108) | -0.020 (0.073) | -0.874** (0.113) | -0.302** (0.110) | -0.299** (0.092) |
| Observations | 4,951 | 4,948 | 4,948 | 4,401 | 4,948 | 4,585 | 4,948 |
| Panel C. Effects of Public and Business R&D | | | | | | | |
| ln(public R&D) _{t-1} | 0.152** (0.043) | 0.158** (0.041) | 0.176** (0.042) | 0.153** (0.045) | 0.188** (0.039) | 0.202** (0.038) | 0.166** (0.041) |
| ln(int. public R&D) _{t-1} | | 0.087 (0.057) | 0.421** (0.160) | 0.068 (0.090) | 0.600** (0.162) | 0.408** (0.134) | 0.131** (0.048) |
| ln(int. business R&D) _{t-1} | | -0.133** (0.062) | -0.339** (0.147) | -0.068 (0.114) | -1.041** (0.136) | -0.545** (0.156) | -0.386** (0.097) |
| Observations | 4,951 | 4,948 | 4,948 | 4,401 | 4,948 | 4,585 | 4,948 |

Notes: ** significant at 5% level, * significant at 10% level. Each column in each panel presents a separate regression. Two-way clustered standard errors at the industry-country and country-year level. All columns use a full set of country*industry fixed effects and industry*year fixed effects. All columns include weights given by the employment share of an industry within a country in the first year it appears in the sample. The dependent variable is private R&D, i.e., R&D conducted in the business sector (BERD) that is also financed by the private sector (i.e., excludes government financed R&D). “Public R&D” is government-financed R&D performed by private firms. “Int. public R&D” is the weighted average of other countries’ public R&D in the same industry and year, where each column uses different weights. “Int. total R&D” is the weighted average of other countries’ public and private R&D in the same industry and year, where each column uses different weights. Weights ω_{kl} are a “distance” measure between country k and country l measured by: (2) GDP per capita. $\omega_{kl} = 1/|GDPcap_k - GDPcap_l|$; (3) Geographic distance. $\omega_{kl} = 1/dist_{kl}$ where geographic distance is measured in kilometers between capital cities; (4) Skill intensity. $\omega_{kl} = 1/|tert_k - tert_l|$ with $tert_l$ being the share of population with tertiary education in country l ; (5) Patent similarity. $\omega_{kl} = \sqrt{\sum_i (pat_{ik} - pat_{il})^2}$, pat_{ik} and pat_{il} are patent share of a specific technology class I (out of 15 patent technology classes) in country k or l (of total patents in that country); (6) Inward FDI flows. $\omega_{kl} = \frac{FDI_{kl}}{\sum_{j=1}^J FDI_{kj}}$ where FDI_{kl} is the amount of FDI that country k receives from country l ; (7) R&D intensity. $\omega_{kl} = 1/\left|\frac{Total\ R\&D_k}{GDP_k} - \frac{Total\ R\&D_l}{GDP_l}\right|$.

APPENDIX TABLE A10: INTERNATIONAL SPILLOVER EFFECTS OF R&D SUBSIDIES ON TFP GROWTH

| Measure of neighbor for spillover | Baseline (1) | Per-capita GDP (2) | Geo- graphy (3) | Skill Intensity (4) | Tech- nology (5) | FDI Flows (6) | R&D Intensity (7) |
|--|--------------------|-----------------------|-----------------------|---------------------------|------------------------|---------------------|-------------------------|
| Dependent variable: TFP Growth | | | | | | | |
| Panel A. OLS | | | | | | | |
| (R&D/value added) _{t-1} | 0.098** (0.040) | 0.054* (0.032) | 0.032 (0.024) | 0.046 (0.033) | 0.020 (0.022) | 0.082** (0.037) | 0.038 (0.029) |
| (International R&D/value added) _{t-1} | | 0.180** (0.053) | 0.212** (0.041) | 0.215** (0.047) | 0.254** (0.049) | 0.073** (0.030) | 0.184** (0.052) |
| Observations | 6,608 | 6,602 | 6,602 | 5,748 | 6,602 | 6,602 | 6,602 |
| Panel B. IV | | | | | | | |
| (R&D/value added) _{t-1} | 0.098** (0.041) | 0.035 (0.044) | 0.036 (0.032) | 0.054 (0.049) | 0.022 (0.033) | 0.043 (0.033) | 0.032 (0.034) |
| (International R&D/value added) _{t-1} | | 0.188** (0.086) | 0.205** (0.055) | 0.183** (0.070) | 0.217** (0.062) | 0.140** (0.040) | 0.184** (0.068) |
| Observations | 6,608 | 5,155 | 5,155 | 4,709 | 5,155 | 5,146 | 5,155 |
| First stage F | | 5.793 | 7.614 | 6.274 | 6.961 | 7.546 | 6.741 |

Notes: ** significant at 5% level, * significant at 10% level. Standard errors are two-way clustered at the industry*country and country*year level. All regressions include a full set of country and year fixed effects. All columns use weights given by the employment share of an industry within a country in the first year it appears in the sample. The dependent variable is TFP growth over 1 year. “International R&D/value added” is the weighted average of other countries’ total R&D/value added in the same industry and year, where each column uses different weights. “R&D” is the total R&D conducted in the business sector (both financed by the private sector and by government). Weights ω_{kl} are a “distance” measure between country k and country l and measured by: (2) GDP per capita. $\omega_{kl} = 1/abs(GDPcap_k - GDPcap_l)$; (3) Geographic distance. $\omega_{kl} = 1/dist_{kl}$ where geographical distance is measured in kms between capital cities; (4) Skill intensity. $\omega_{kl} = 1/abs(tert_k - tert_l)$ with $tert_l$ being the share of population with tertiary education in country l ; (5) Patent similarity. $\omega_{kl} = \sqrt{\sum_i (pat_{ik} - pat_{il})^2}$, pat_{ik} and pat_{il} are patent share of a specific technology class i (out of 15 patent technology classes) in country k or l (of total patents in that country); (6) Inward FDI flows. $\omega_{kl} = \frac{FDI_{kl}}{\sum_{j=1}^J FDI_{kj}}$ where FDI_{kl} is the amount of FDI that country k receives from country l ; (7) R&D intensity. $\omega_{kl} = 1/abs\left(\frac{Total\ R\&D_k}{GDP_k} - \frac{Total\ R\&D_l}{GDP_l}\right)$

**APPENDIX TABLE A11: INTERNATIONAL DISPLACEMENT EFFECTS OF PUBLIC R&D
– FOR ILLUSTRATION OF MAGNITUDE OF EFFECTS**

| Measure of neighbor for spillover | Baseline | Per-capita GDP (1) | Geography (2) | Skill Intensity (3) | Tech-nology (4) | FDI Flows (5) | R&D Intensity (6) | Trade (imports) (7) | Trade (imports) (8) |
|--|--------------------|-----------------------|---------------------|------------------------|---------------------|--------------------|----------------------|------------------------|------------------------|
| Dependent variable: ln(Privately funded business R&D) | | | | | | | | | |
| ln(Domestic Public R&D) _t | 0.235** (0.038) | 0.243** (0.038) | 0.242** (0.038) | 0.247** (0.040) | 0.239** (0.037) | 0.251** (0.037) | 0.244** (0.037) | 0.246** (0.037) | 0.244** (0.038) |
| ln(Int. Total R&D) _{t-1} | | 0.001 (0.064) | -0.325** (0.102) | -0.086 (0.086) | -0.850** (0.277) | -0.086 (0.077) | -0.278** (0.086) | -0.219** (0.070) | -0.181** (0.077) |
| ln(output) _t | 0.674** (0.172) | 0.637** (0.179) | 0.574** (0.180) | 0.628** (0.183) | 0.557** (0.175) | 0.637** (0.180) | 0.636** (0.180) | 0.610** (0.181) | 0.613** (0.178) |
| Observations | 6,671 | 6,479 | 6,479 | 5,674 | 6,479 | 6,443 | 6,479 | 6,460 | 6,479 |

Notes: ** significant at 5% level, * significant at 10% level. Each column in each panel constitutes a separate regression. Two-way clustered standard errors at the industry*country and country*year level. All columns include full sets of country by industry fixed effects, industry by year fixed effects, and country by year fixed effects. All columns use weights given by the employment share of an industry within a country in the first year it appears in the sample. The dependent variable is private R&D, i.e. R&D conducted in the business sector (BERD) that is also financed by the private sector (i.e. excludes government financed -R&D). “Public R&D” is government financed R&D performed by private firms. “Int. Total R&D” is the weighted average of other countries’ public and private R&D in the same industry and year, where each column uses different weights. Weights ω_{kl} are a “distance” measure between country k and country l and measured by: (1) GDP per capita. $\omega_{kl} = 1/abs(GDPcap_k - GDPcap_l)$; (2) Geographic distance. $\omega_{kl} = 1/dist_{kl}$ where geographical distance is measured in kms between capital cities; (3) Skill intensity. $\omega_{kl} = 1/abs(tert_k - tert_l)$ with $tert_l$ being the share of population with tertiary education in country l ; (4) Patent similarity. $\omega_{kl} = \sqrt{\sum_i (pat_{ik} - pat_{il})^2}$, pat_{ik} and pat_{il} are patent share of a specific technology class i (out of 15 patent technology classes) in country k or l (of total patents in that country); (5) Inward FDI flows. $\omega_{kl} = \frac{FDI_{kl}}{\sum_{j=1}^J FDI_{kj}}$ where FDI_{kl} is the amount of FDI that country k receives from country l ; (6) R&D intensity. $\omega_{kl} = 1/abs\left(\frac{Total\ R\&D_k}{GDP_k} - \frac{Total\ R\&D_l}{GDP_l}\right)$; (7) Import share. $\omega_{kl} = \frac{IMP_{kl}}{\sum_{j=1}^J IMP_{kj}}$ where IMP_{kl} are imports from country l to country k ; (8) Export share. $\omega_{kl} = \frac{EXP_{kl}}{\sum_{j=1}^J EXP_{kj}}$ where EXP_{kl} are exports from country l to country k .