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Impact of exchange rate move-ments on exports: An analysis of Indian non-financial sector firms

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Yin-Wong Cheung and Rajeswari Sengupta

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Contents

Abstract.....	4
1 Introduction.....	5
2 Data.....	9
3 Baseline regression.....	12
4 Firm-specific effects.....	15
5 Macro effects.....	17
6 Additional analyses.....	19
7 Concluding remarks.....	25
References.....	28
Data appendix.....	30

Yin-Wong Cheung and Rajeswari Sengupta

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Abstract

We explore the real effective exchange rate (REER) effects on the share of exports of Indian non-financial sector firms for the period 2000 to 2010. Our empirical analysis reveals that, on average, there has been a strong and significant negative impact from currency appreciation and currency volatility on market shares of India's exporting firms. Labor costs are found to amplify the exchange-rate effects on trade. Further, there is evidence that the Indian firms considered here respond asymmetrically to exchange rates. A REER change effect, for example, is more likely to arise from a negative appreciation effect than a depreciation effect. Indian firms with smaller export shares tend to respond more strongly to both REER change and volatility than those with larger export shares. Services exporters are impacted more strongly by exchange rate fluctuations than firms exporting goods. The findings on asymmetric responses, in particular, have important policy implications.

JEL classifications: F1, F4

Keywords: exchange rate fluctuations, firm-level export shares, asymmetric effects, services exports

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1 Introduction

The share of global trade in total world output has almost tripled since the World War II. In the last quarter century, the global economy witnessed not just a rapid expansion in international trade but also the growing prominence of dynamic emerging economies on the global trade landscape. Emerging markets have steadily become important trading centers thanks to the growing role of global supply chains and high-technology exports.

Despite steady growth in global trade, however, there are recurring concerns about the impact of exchange-rate movements on trade in general and on a country's export and import activities. The collapse of the Bretton Woods system in the early 1970s triggered a wave of debate on whether exchange-rate variability hampered global trade. The discussion of exchange-rate effects on trade was rekindled more recently by the 1997 Asian financial crisis and the 2008 global financial crisis.

The overall trade activity of a country is an aggregation of decisions of individual firms. Hence, understanding the effects of exchange-rate changes on the trade balance calls for an analysis of how exchange-rate fluctuations affect the decisions of individual firms. Such analysis provides insights into heterogeneous responses across firms to exchange-rate movements and the related policy implications of central bank efforts to manage and stabilize foreign-exchange fluctuations.

India provides an interesting case study in how exchange-rate fluctuations impact exports. During the 1960s and 1970s, India was one of the world's least open economies. Indeed, India's exchange rate was more or less fixed until 1991, when the country embraced a reform agenda and implemented a host of liberalization measures that primarily targeted the foreign-exchange market and tradable sectors. India adopted a more market-oriented exchange rate regime in the first half of 1990s. Since deregulation, the rupee's exchange-rate has mostly been in a managed floating regime under which the Reserve Bank of India intervenes from time to time to stabilize the nominal exchange rate.¹

The current study focuses on India's exports. Compared with other firms, exporting firms are usually associated with higher levels of productivity and profitability. Moreover, a strong export sector might generate spillovers to other sectors that promote overall economic growth. India has witnessed strong economic performance coupled with a strong

¹ The Reserve Bank of India appears to have adopted an asymmetric intervention policy designed to restrain currency appreciation while allowing reasonable amounts of depreciation (Sen Gupta and Sengupta, 2013).

export sector in the past decade or so. Thus, it is quite conceivable that policies that promote exports are conducive to economic growth.

The annual growth rate of India's exports of goods and services increased from 16% in 1999–2000 to around 33% in 2010–2011. The contribution of exports to GDP went up from 6% in 1990 to 12% in 2000 and 23% in 2010. Simultaneously, India's overall share of total world trade (which includes trade in both goods and services) increased from 0.5% in 1990 to about 1.4% in 2010. As a result, India moved up seven places between 1999 and 2009, securing a ranking of the world's 14th largest trading center.

During the period 2000–2010, growth in exports of commercial services (an average of 23% p.a.) outstripped merchandise exports (18% p.a.). It is striking that the high export growth occurred despite appreciation of about 1.4% in India's real effective exchange rate (REER) over the period.

India's REER has mostly appreciated since 1994. Anecdotal evidence suggests that up to 1993, the relationship of REER to total exports followed the textbook prescription: i.e. exchange-rate depreciation positively affected exports. Since 1993, the trend has defied traditional expectations. However, as pointed out by Veeramani (2008), the fact that Indian exports grew rapidly after 2000 despite REER appreciation does not necessarily imply that REER appreciation had no adverse impact on exports. The growth rate of exports might have been higher without REER appreciation.

Against this background, we ask how fluctuations in the exchange rate have affected decisions of Indian exporting firms and consider whether the data suggests a weakening of the link between REER and exports. Owing to data restrictions, we focus on export behavior in the decade of the 2000s.

A few studies, notably Veeramani (2008) and Srinivasan and Wallack (2003), consider the impact of exchange-rate changes on overall exports in the Indian context. However, no study to date has used micro-evidence to explore the issue. Accordingly, our objective in this paper is to use detailed firm-level data from a sample of Indian non-financial sector firms to investigate empirically the exchange-rate effect on the exporting behavior of firms, controlling for other possible determinants. Beside exchange-rate changes, we investigate the implication of exchange-rate volatility for trade. The trade effect of exchange rate volatility has been an intensely debated issue since the breakdown of the Bretton Woods system. Interestingly, neither the theoretical nor the empirical studies

offer any firm conclusions on the effect of exchange-rate volatility on international trade flows (Côté, 2004; Cheung, 2005).

Numerous studies in the empirical trade literature explore the aggregate relationships between the exchange rate and international trade at the country level. However, studies using aggregate data are subject to problems such as aggregation bias (Dekle *et al.*, 2007), simultaneity (Adolfson, 2001), and measurement error in constructing aggregate indices. There is relatively little empirical work on the responses of exports at the level of firms or individual producers. Exceptions include Fitzgerald and Haller (2010), Berman, Martin and Meyer (2009), Greenaway, Kneller, and Zhang (2007), Campa (2004), Bernard and Jensen (2004a, b), Bugamelli and Infante (2003), and Forbes (2002).

Most empirical studies using micro-evidence consider developed countries. The micro-data evidence on the impact of exchange-rate movements on the decisions of individual producers to export is ambiguous. All these studies find quantitatively small effects of exchange-rate movements on entry and exit, as well as changes in exports, due to the fact that exchange-rate movements come mainly from existing exporters adjusting their production (intensive margin) rather than new entrants (extensive margin).

Our study focuses on exchange-rate movements and exports for a large panel of Indian firms. Arguably, it is the first extensive firm-level study on India's firm exporting behavior. Thus, this paper contributes to the growing literature on the responses of individual firms to exchange-rate variation and on trade behavior in developing countries. The use of a rich firm-level data set enables us to exploit heterogeneity across firms in a large developing economy and alleviate biases such as those arising from aggregation. Given the reported weak extensive margin effect, we investigate mainly the intensive margin effect.²

The questions we are primarily interested in are:

- What is the impact of exchange-rate depreciation (appreciation) on exports of Indian manufacturing firms?
- Does the textbook prediction that exchange-rate depreciation (appreciation) boosts (deters) exports hold for Indian firms or is there no significant association?
- What are the firm-specific features that influence their export responses to exchange rate changes?
- What are the macro features of the economy as a whole that impact firm-level export responses to exchange-rate movements?

² Our future research agenda includes analyses of the extensive margin and the impact of exchange-rate changes on the entry and exit of Indian firms to and from the export market.

To anticipate results, our baseline empirical analysis reveals that, for our generic exports equation that uses exchange-rate and income as explanatory variables, a one percentage point increase in the appreciation of the REER causes a 6.3% reduction in the change of the share of exports in total sales of Indian firms. In the presence of certain control variables, however, this exchange-rate effect can exceed 10%. The result is largely in line with other relevant studies, including Virmani (1991), Joshi and Little (1994), Srinivasan (1998), Srinivasan and Wallack (2003) and Veermani (2007, 2008) that use aggregate data to demonstrate the negative REER appreciation effect on India's aggregate merchandise exports. Our empirical findings in general are also suggestive of a negative volatility effect on the export shares of firms. A one standard deviation decline in REER volatility on average increases an Indian firm's export share by as much as 13%. Both results are in line with the predictions of standard theory.

In addition to the baseline formulation, we consider alternative specifications to evaluate the exchange-rate effects. Notable findings include:

- For Indian firms, firm-specific accounting information does not seem to affect the exchange-rate and trade interaction.
- Both real and nominal wage increases have a negative effect on exports.
- Appreciation is associated with a stronger exchange-rate change and a stronger volatility effect on trade than depreciation.
- Exchange-rate effects are different for firms with large and small export shares.
- Exchange-rate changes have a stronger impact on services exports than merchandise exports.
- A negative exchange-rate change effect is found for most of our additional analyses.

In the next section, we describe the data set used for our analysis. The baseline regression is reported in section 3. The analyses based on alternative specifications are reported in sections 4 to 6. Some concluding remarks are offered in section 7.

2 Data

Our firm-level data are drawn from the Prowess database of the Center for Monitoring Indian Economy (CMIE). The database provides annual information on publicly traded non-financial firms (both consolidated and stand-alone). Here, we focus on exporting firms between 2000 and 2010, the period for which data is available.

Table 1 Number of exporting firms, 2000–2010

Year	2000	2001	2002	2003	2004	2005
Total exporting firms	3214	3251	3698	3348	4105	4154
Year	2006	2007	2008	2009	2010	
Total exporting firms	4167	4225	4289	4068	3702	

Tables 1 to 3 present some descriptive statistics of our firm-level data. These are only for firms with positive exports, i.e. firms that exported at least once during the sample period. The number of firms in each of the sample years, after dropping out outliers, is listed in Table 1.³ The size of our firm-sample grows monotonically from 3,214 firms in 2000 to 4,289 firms in 2008. Following the 2008 global finance crisis, it drops to 4,068 and then to 3,702 in 2010. Approximately 18% of the firms are present for all eleven years in our observation period, close to 13% firms are present in the sample for at least 8 years, 56% are observed for at least 3 years and only 8% make one-year appearances.

Some summary statistics for the firm-specific and macro variables are given in Tables 2 and 3. Table 2 covers the entire sample period and the firm-specific summary statistics are based on the full firm-sample. Table 3, for comparison purposes, shows the summary statistics for three selected years: 2000, 2005, and 2010.

³ Extensive checks were conducted in preparing the sample. We drop all firm/year observations if the accounting data are not self-consistent. In particular, we drop observations if firm-level accounting variables do not accord with sign conventions (e.g. if sales, total assets, or exports are negative, if exports exceed 100% of sales, if foreign currency borrowing exceeds total liabilities). Firms with zero sales are also excluded.

Table 2 Descriptive statistics of some important variables: full sample

Variables (%)	Mean	Std. Dev.	Min	Max
Exports/sales	28.03	31.91	0	100.99
Exports/sales (goods)	25.07	29.60	0	100.99
Exports/sales (services)	26.19	35.51	0	100.98
Firm size (log of total assets)	6.64	1.83	0	14.86
Capacity utilization	110.27	75.74	0	599.37
Collateral (net fixed assets/total assets)	32.30	20.84	0	100
Foreign liability/total liability	1.66	6.08	0	99.46
REER	93.35	3.44	89.52	100
REER change	1.35	4.88	-5.45	11.71
REER volatility	1.90	0.69	1.22	3.54
Nominal wage index	219.04	127.46	87.40	418.1
Real per capita GDP growth rate	5.54	2.37	2.23	8.22
World exports/world GDP	26.30	2.03	24.02	29.57

The average firm in our sample exports around 28% of its sales. This percentage goes up from 26% to almost 29% between 2000 and 2005, then declines to 27% in 2010. Exports of goods and services display different patterns. While share of exports of goods in total sales does not exhibit much fluctuation across years and hovers around an average of 25%, services exports increase from 25% to 31% between 2000 and 2005 and remain there through 2010. Across all firms, the exports/sales ratio appears quite variable and has a standard deviation of around 32%. Further, the exports/sales ratio has a larger degree of dispersion for exports of services than for exports of goods. The variability allows us to discriminate between different behavioral patterns across firms.

Table 3 Descriptive statistics for select variables in 2000, 2005, and 2010

Variables (%)	Mean	Std. Dev.	Mean	Std. Dev.	Mean	Std. Dev.
	2000		2005		2010	
Exports/sales	26.18	31.27	28.89	32.31	27.04	31.22
Exports/sales (goods)	25.92	50.13	26.15	35.72	23.99	28.77
Exports/sales (services)	25.59	53.83	31.29	55.66	30.54	73.35
Firm size (log of total assets)	6.32	1.63	6.45	1.81	7.29	1.93
Capacity utilization (sales/total assets)	103.81	72.56	114.38	75.43	106.88	72.44
Collateral (net fixed assets/total assets)	36.28	21.41	31.52	20.41	29.60	20.16
Foreign liability/total liability	0.73	4.35	2.15	6.80	1.75	5.64
REER	91.34	0	94.10	0	100.001	0
REER change	4.53	0	3.40	0	11.71	0
REER volatility	1.91	0	1.74	0	2.19	0
Nominal wage index	103.76	0	100	0	418.1	0
Real per capita GDP growth	2.23	0	7.65	0	8.04	0
World exports/world GDP	24.70	0	26.77	0	27.97	0

Other firm-level characteristics such as capacity utilization and share of foreign-currency borrowing seem to follow the pattern of exports/sales ratio, i.e. the 2005 values are higher than 2000 and 2010 values. Average firm size tends to increase over time, while collateral appears to decline over time.

Our main explanatory variable, REER, exhibits steady appreciation from 91 to 100 between 2000 and 2010, and registers an appreciation of nearly 12% in 2010.⁴ Compared with data for 2000 and 2005, both REER volatility and wages registered large values

⁴ Several studies use destination-specific bilateral real or nominal exchange-rates. Unfortunately, destination-specific information is not available for Indian firms. Given our data structure and limitations, aggregate REER and NEER were the only options available. We also lacked data on export volumes, and hence had to work with exports/sales ratio based on data values. In section 6, as a part of our robustness checks, we consider WPI- and CPI- based real exports data.

in 2010. Despite the commonly perceived adverse effect of the global financial crisis on trade volume, the world exports/GDP ratio in 2010 is slightly higher than the levels in 2000 and 2005. Details on the macroeconomic variables used in the analysis are provided in the Data Appendix.

3 Baseline regression

Our baseline specification for studying exchange-rate effects is given by:

$$\Delta X_{it} = \beta_0 + \beta_1 \Delta REER_t + \beta_2 (REER_vol)_t + \beta_3 \Delta I_t^* + \chi_i \mu_i + \phi \eta_t + \varepsilon_{it} \quad (1)$$

where X_{it} represents firm-level export shares defined as the exports to sales ratio of firm i at time t ; $\Delta REER_t$ is the change in real effective exchange rate (REER) with an increase indicating rupee appreciation; ΔI_t^* is change in the level of foreign income (measured by trade-share-weighted average of incomes of India's top five trading partners) that represents the general growth in overseas markets; and $(REER_vol)_t$ is the volatility of real effective exchange rate measured using standard deviation of monthly REER indices of the year. Both contemporaneous $\Delta REER_t$ and $(REER_vol)_t$ are used as it is unlikely that a firm's exporting behavior will have an effect on either REER change or REER volatility.⁵ Thus, Equation (1) assesses the average responses in changes in the export shares of firms to exchange-rate change and variability. The exchange-rate variables and income variable are the explanatory factors in a canonical exports demand equation.

A crucial problem for firm-level studies is the classical omitted variable problem caused by unobserved firm characteristics. One solution is to control for as many firm-level variables as possible. Here, however, there is an obvious limitation imposed by the data set. Thus, a fixed-effects variable, μ_i , is included to capture a firm's specific attributes that are (approximately) time-invariant and have implications for exporting behavior. These attributes may include managerial characteristics, foreign experience, and product

⁵ In fact, the inclusion of the lagged REER change and lagged REER volatility in the canonical regression does not change the significant negative exchange rate results. These results are available upon request.

quality.⁶ In the next section, we introduce an augmented specification that incorporates time-varying firm-specific explanatory variables.

The term η_t is an annual time effect reflecting temporal variations in export shares that are common to all firms in the sample. These year fixed effects control for common national-level or macro shocks such as changes in the business cycle, trade liberalization across all firms, general technology advances, as well as global business cycle effects. Finally, ε_{it} is the regression error term.

Table 4 Baseline regression I with change in exports/sales as dependent variable

Variables	I	II	III	IV	V
REER change	-0.063*** (0.017)		-0.062*** (0.017)	-0.063*** (0.017)	-0.071*** (0.019)
REER volatility		-0.379*** (0.119)	-0.375*** (0.119)	-0.455*** (0.120)	-0.303* (0.174)
I* change				0.027*** (0.009)	0.033*** (0.010)
Year dummies	no	no	no	no	yes
Firm fixed effects	yes	yes	yes	yes	yes
Observations	33132	33132	33132	33132	33132

Note: Robust Standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1, Constant is not reported. The table presents results of estimating Equation (1). I^* is the level of foreign income (proxied by trade-share-weighted average of incomes of India's principal trading partners) that represents growth in overseas markets.

The results from estimating Equation (1) are reported in Table 4. For the sake of brevity, the time and fixed effects variables (η_t and μ_i) included in the regression are not reported. Columns I and II present the individual effects of REER change and volatility on firms' export shares. The joint effects appear in Column III.

Controlling for fixed effects, both exchange-rate movements and volatility have a statistically significant negative impact on export shares. More precisely, a one percentage

⁶ We incorporated sector specific or industry dummy variables to capture characteristics that are specific to a firm's main sector. The results were found to be robust to the presence of these dummy variables. They are not reported here for the sake of brevity, but are available on request.

point appreciation of the REER reduces the average firm's export share by 6.3%. A one standard-deviation rise in REER volatility (or 0.69 as in Table 2) dampens exports by as much as 26% (0.69×0.379).

The significant exchange-rate effect accords with result predicted by standard economic theory, i.e. a higher price level hurts exports. The effect also echoes Srinivasan and Wallack (2003) and Veeramani (2008), who report a negative relationship between the real exchange rate and merchandise aggregate exports in India.⁷ This negative volatility effect lends support to the reasoning that a high level of uncertainty represented by a high level of volatility has an adverse effect on trade.⁸ The results in Column III suggest the overlap between the two exchange-rate effects on exports is rather limited. The two coefficient estimates under Column III are quite comparable to their corresponding estimates in Columns I and II.

The marginal effect of foreign income on export shares appears in Columns IV and V. The positive and significant income effect is in line with standard textbook predictions and is consistent with Srinivasan and Wallack (2003), Bugamelli and Infante (2003) and Veeramani (2008). It may be noted here that when the specification includes the time effect variables, the REER effect is strengthened (e.g. from 6.3% to 7.1%).

In summary, the exchange-rate effects based on India's firm-level data are largely in line with standard trade theory. These results are also consistent with the findings of similar studies on the topic that are based on micro-data from developed economies such as Berman, Martin, and Meyer (2009), Bernard and Jensen (2004b), Campa (2004), Fitzgerald and Haller (2010), Forbes (2002), and Greenaway, Kneller, and Zhang (2007).

⁷ Negative exchange-rate effects based on firm-level data for developed countries are reported e.g. in Greenaway, Kneller, and Zhang (2007), Forbes (2002), and Bernard and Jensen (2004b).

⁸ See Clark (1973), Baron (1976), Hooper and Kohlhagen (1978) among some of the earlier evidence of the negative impact of exchange rate volatility on trade and Côte (1994) for an extensive literature review on this effect.

4 Firm-specific effects

To control for effects of firm-specific variables, we consider the regression specification:

$$\Delta X_{it} = \beta_0 + \beta_1 \Delta REER_t + \beta_2 (REER_vol)_t + \beta_3 \Delta I_t^* + \beta_4 Y_{it-1} + \chi \mu_i + \phi \eta_t + \varepsilon_{it}, \quad (2)$$

where Y_{it-1} is a vector of observable characteristics of firm i at time $t-1$, that could affect a firm's export decision. The time-varying firm-level explanatory variables are all lagged by one year to avoid any simultaneity effect. We also include firm-specific variables in their first differences. The results are qualitatively similar, and hence are not reported here for the sake of brevity.

The firm-specific variables included are:

- firm size measured by the log of total assets (this is often interpreted as a proxy for firm success or ability to cope with financial constraints);
- firm efficiency in capital utilization measured as the ratio of sales to total assets, (i.e. more efficient firms are likely to handle unfavorable exchange-rate movements better);
- a firm's ability to borrow externally captured by a measure of collateral (e.g. ratio of net fixed assets to total assets); and
- a firm's foreign-currency borrowing measured by the ratio of secured and unsecured foreign-currency borrowing to total liabilities.

While currency depreciation increases the local currency burden of foreign currency debt thereby adversely affecting a firm's balance sheet, exporters have a natural hedge against currency depreciation by virtue of their export revenues denominated in foreign currency. Thus, the extent of foreign-currency borrowing may have an overall ambiguous implication for exchange-rate effects on exports. These firm characteristics have been used in exploring factors determining exports of firms in studies such as Roberts and Tybout (1997), Campa (2004), Greenaway, Guariglia, and Kneller (2005), and Bernard and Jensen (2004b).

The results of estimating Equation (2) are reported in Table 5. We sequentially introduce the firm-level determinants to the baseline regression. A surprising finding is that none of these firm-level variables has a significant effect on the export shares of firms. The insignificant result is quite different from those studies based on data from developed

countries (Greenaway, Kneller, and Zhang, 2007; Bugamelli and Infante, 2003 and Forbes, 2002).

We further note that some efforts were made to ensure that the adopted specification is not the main reason driving the insignificant results, including the one pertaining to firm size. Specifically, we considered a specification that included contemporaneous, instead of lagged, firm-level variables. On the firm size variable, we used the log of total number of employees to define firm size, an alternative measure of firm size employed by some studies. We also interacted the REER change and volatility variables with the firm size measure. Results derived from these modifications are qualitatively similar to those in Table 5 and thus are not reported here for the sake of brevity.

Table 5 Estimating equation 2 with lagged firm-level controls added

Variables	I	II	III	IV
REER change	-0.070*** (0.020)	-0.069*** (0.020)	-0.072*** (0.020)	-0.073*** (0.020)
REER volatility	-0.293 (0.183)	-0.285 (0.183)	-0.312* (0.185)	-0.322* (0.184)
I* change	0.033*** (0.011)	0.034*** (0.011)	0.037*** (0.011)	0.036*** (0.011)
Lagged firm size	-0.038 (0.209)	-0.042 (0.213)	0.090 (0.236)	0.127 (0.234)
Lagged collateral		0.008 (0.010)	0.009 (0.010)	0.010 (0.010)
Lagged capacity utilization			0.002 (0.003)	0.002 (0.003)
Lagged foreign liability share				-0.026 (0.017)
Firm fixed effects	yes	yes	yes	yes
Year dummies	yes	yes	yes	yes
Observations	33132	33021	32930	32922

Note: Robust Standard errors in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Constant is not reported. The table presents results of estimating Equation (2). *Firm size* is measured using the logarithm of total assets. *Collateral* is defined as the ratio of net fixed assets to total assets, *capacity utilization* is measured by the ratio of sales to total assets, and *foreign liability share* is the ratio of foreign currency liabilities (secured and unsecured foreign currency borrowing) and total liabilities. All firm-level control variables are lagged by one year.

The inclusion of these firm-specific variables in general has limited impact on the exchange-rate and income variables. The REER volatility variable becomes marginally insignificant at the 10% level in the presence of firm size and collateral variables. However, the results from other specifications do not confirm the insignificance of the volatility variable. Indeed, the inclusion of insignificant firm-level variables might have impaired the regression efficiency and reduced the level of significance of other variables.

5 Macro effects

We estimate the effect of India's macro conditions on its firms' export decision using the following specification:

$$\Delta X_{it} = \beta_0 + \beta_1 \Delta REER_t + \beta_2 (REER_vol)_t + \beta_3 \Delta I_t^* + \beta_4 Y_{it-1} + \beta_5 Z_t + \chi \mu_i + \gamma \eta_i + \varepsilon_{it}, \quad (3)$$

where the term Z_t includes (i) percentage change in nominal (or real) wage index, (ii) percentage change in world exports to GDP ratio, and (iii) real per capita GDP growth rate. The wage variables are included to capture operation costs, the world exports to GDP ratio to assess the global trade effect, and GDP growth rate to reflect aggregate domestic demand.

From the point of view of an average firm, it is unlikely that its exporting behavior will have a noticeable impact on either REER change or REER volatility or any of the other macro control variables. Thus, we incorporate the contemporaneous, un-lagged values of the macro variables, which represent common shocks to all firms. All these macro variables have been examined in other empirical trade studies, including Bernard and Jensen (2004a), Greenaway *et al.* (2007), and Bugamelli and Infante (2003).⁹

Table 6 shows the estimation results after incorporating macro-level factors. Since the firm-specific accounting variables considered in Section 4 are insignificant when included in the regression, they are not reported or discussed here for the sake of brevity. Adding or dropping these firm-specific variables does not qualitatively influence the effects of the macro-variables.

⁹ As mentioned earlier, our baseline results are in general robust to the inclusion of lagged REER change and lagged REER volatility variables.

Nominal wages have a negative effect on firms' export share (Column I). The result is quite intuitive; a rise in wages increases operation costs that then reduce the firm's competitiveness in the global market. A similar negative effect is also obtained when real wages are used (Column II). This is consistent with other related studies such as Bugamelli and Infante (2003), who find for their sample of Italian firms that the probability of exporting decreases with average wages (with wages here as a proxy for cost competitiveness).

Apparently, exports of Indian firms co-move with world exports. The world exports to GDP ratio that reflects the general trend in global trade, is found to have a significantly positive coefficient estimate (Column III). Our results suggest that the global trade pattern points to the general behavior of Indian exporting firms, but does not overshadow the exchange-rate effect. Indeed, the coefficient estimate of the exchange-rate effect is largest under Column III.

Table 6 Estimating equation 3 with country-level controls added

Variables	I	II	III	IV
REER change	-0.100*** (0.023)	-0.103*** (0.024)	-0.172*** (0.039)	-0.108*** (0.043)
REER volatility	-0.938*** (0.335)	-0.959*** (0.345)	-0.700*** (0.280)	-0.663*** (0.285)
I* change	0.066*** (0.019)	0.067*** (0.004)		
Change in nominal wages	-0.006* (0.003)		-0.004 (0.003)	-0.001 (0.003)
Change in real wages		-0.007*** (0.004)		
Change in world exports to GDP ratio			0.107*** (0.032)	0.069** (0.033)
Real per capita GDP growth				-0.131*** (0.053)
Firm fixed effects	yes	yes	yes	yes
Year dummies	yes	yes	yes	yes
Observations	33132	33132	33132	33132

Note: Robust standard errors in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Constant is not reported. The table presents results of estimating Equation (3). *Change in nominal wages* denotes percentage change in hourly wages in local currency over previous year based on data from the Yearbook of Labor Statistics. *Change in real wages* denotes percentage change in hourly wages in local currency adjusted for inflation over previous year based on data from the Yearbook of Labor Statistics. *Real per capita GDP growth rate* is the annual percentage change in per capita GDP (constant local currency units) based on data from the WDI database.

The coefficient estimate of real per capita GDP growth rate is negative, indicating the tendency that export activity declines with an increase in domestic demand that is proxied by GDP growth (Column III). The negative effect of domestic demand is similar to that found for Italian firms in Bugamelli and Infante (2003).

Notably, the inclusion of these macro variables renders the wage effect insignificant, while amplifying the exchange-rate effect. Comparing the corresponding coefficient estimates in Table 4 and Table 6, the effects of REER change and REER volatility appear stronger in Table 6 than in Table 4. The income variable also displays a substantial increase in its impact on export shares in Table 6. Thus, developments in the domestic and global markets might affect the response of firms' export shares to exchange rate and foreign income.

6 Additional analyses

To deepen our investigation, we conduct a few additional analyses. First, we consider asymmetric exchange-rate change effects by constructing dummy variables for appreciation and depreciation and interact them with the REER change and volatility terms. The results allowing for exchange-rate asymmetry are reported in Columns I and II of Table 7. Apparently, the exchange-rate effects observed in previous sections are driven by exchange-rate appreciation but not depreciation.

The coefficient estimate of REER change interacted with the appreciation dummy is statistically significant with a negative sign, implying, as expected, that a large appreciation reduces exports. The magnitude of the estimate is larger in absolute value than the corresponding estimate in Table 4. However, when interacted with the depreciation dummy, the REER change variable is not significant. The REER volatility variable displays similar asymmetric effects. During the appreciation phase, REER volatility has a strong and statistically significant negative effect on export shares. Its effect is insignificant when the REER is depreciating. In other words, appreciation hurts export activity.

Next, we assess the dependence of response to exchange-rates on a firm's level of exports. We construct small and large exports dummy variables based on whether the export shares are above or below the median level. We then interact these dummy variables with the REER change and volatility terms. Results are shown in Columns III and IV of Table 7.

Relatively speaking, the adverse exchange-rate effect seems to be stronger for firms with a small export share than with a large one. The estimated results indicate that a one percentage point REER appreciation reduces the export share by around 11% for firms with below the median export shares, and by 5% for firms with export shares above the median level. Firms that export relatively less are apparently more adversely affected by appreciation.

Table 7 Additional analysis I: Asymmetric effects of REER change and REER volatility

Variables	I	II	III	IV
I* Change	0.013 (0.012)	0.018* (0.011)	0.025*** (0.010)	-0.042** (0.021)
REER change * Appreciation dummy	-0.163*** (0.053)			
REER change * Depreciation dummy	0.192 (0.126)			
REER volatility * Appreciation dummy		-0.644*** (0.214)		
REER volatility * Depreciation dummy		-0.154 (0.171)		
REER change * Small exports dummy			-0.111*** (0.046)	
REER change * Large exports dummy			-0.053*** (0.019)	
REER volatility * Small exports dummy				-0.930** (0.412)
REER volatility * Large exports dummy				1.263*** (0.413)
Firm fixed effects	yes	yes	yes	yes
Year dummies	yes	yes	yes	yes
Observations	33132	33132	33132	33132

Note: Robust Standard errors in parentheses; *** p<0.01, ** p<0.05, * p<0.1. Constant is not reported. The table presents results of estimating Equation (1). REER change and REER volatility are each interacted with an appreciation dummy variable that takes the value 1 if REER change is positive, and zero otherwise, and with a depreciation dummy variable that takes the value 1 if REER change is negative, and zero otherwise. REER change and REER volatility are also interacted with a small exports dummy that is 1 if the value of exports is less than the median, and 0 otherwise, as well as a large exports dummy that is 1 if the value of exports is greater than or equal to the median value, and 0 otherwise.

On the volatility side, firms with below-median export shares react negatively to REER volatility, while somewhat curiously, firms with larger export shares react positively to a rise in exchange-rate volatility. As mentioned earlier, existing theoretical models and empirical results do not offer a definite verdict on the volatility effect on trade. Although the negative volatility effect appears intuitive and is supported by models based on risk aversion, positive volatility is a possibility in models that are based on transaction cost considerations and that view exporting activity as an option exercised under favorable conditions.¹⁰ By splitting the sample according to the export share of the firm, both positive and negative effects are revealed. Arguably, firms that have a large export share could have the incentive and, possibly, the means to benefit from exchange-rate volatility through hedging or re-directing their exports to alternative markets.

Third, recognizing the growing importance of services trade, we split our firm-sample into those that export goods and those that export services in order to investigate whether the exchange-rate has differential impacts on these two export categories. India's export sector has been dominated by commercial services over the last decade. Table 3 shows that the share of services exports in sales of firms averages around 30%, while the average share of goods exports is only 23% in our sample. Information technology (IT) services are a main component of India's commercial services exports (average export share for the IT services category for our sample period is around 64%), and so we further distinguish between exports of non-IT services and IT services.

The results presented in Table 8 suggest differential exchange-rate effects across export types. Goods exports appear less sensitive to the negative exchange-rate effect than services exports. The coefficient estimates indicate that the effect of REER change on services exports is about 50% stronger than on goods exports.¹¹ Interestingly, while the exports of IT services yield a large (in magnitude) REER change coefficient estimate, the estimate is statistically insignificant. Instead, non-IT services are the only type of services exports that seem to be significantly impacted by exchange-rate change. Thus, exchange-rate management could have different implications for different types of exporting activities. The breakdown of firms into different categories reveals that the foreign income effect

¹⁰ Clark (1973) is an early paper on the negative volatility effect; Franke (1991) offers an example of the positive volatility effect.

¹¹ This result is consistent with Eichengreen and Gupta (2012), who find that the effect of the real exchange rate at the macro level is stronger for exports of services than exports of goods.

is only observed for exports of goods; exports of services are not significantly influenced by income.

Table 8 Additional analysis II: Splitting sample according to export of goods and services

Variables	Exports of goods (I)	Exports of services (II)	Exports of IT services (III)	Exports of non-IT services (IV)
REER change	-0.062*** (0.017)	-0.105** (0.049)	-0.195 (0.177)	-0.068*** (0.021)
REER volatility	-0.198 (0.210)	-0.343 (0.360)	-0.539* (0.315)	-0.425 (0.423)
I* change	0.042*** (0.012)	0.029 (0.018)	-0.031 (0.040)	0.048*** (0.022)
Firm fixed effects	yes	yes	yes	yes
Year dummies	yes	yes	yes	yes
Observations	27073	9247	2456	6928

Note: Robust Standard errors in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Constant is not reported. The dependent variable in Column I is change in the ratio of goods exports/sales. In Column II, it is change in the ratio of services exports/sales. Column III estimates the effect of REER change and volatility on the change in services exports/sales ratio exclusively for IT service industries and Column IV for non-IT service industries.

Fourth, to study the sensitivity of the results to our choices for trade and exchange-rate variables, we consider real exports as the dependent variable, as well as effects of the nominal effective exchange rate (NEER).

Table 9 reports the results of estimating Equation (1) with the exports variable measured by a firm's exports normalized by wholesale or consumer price indices. Normalization using wholesale or consumer price indices does not change the estimation results. Both the exchange-rate and foreign-income effects are qualitatively similar to those reported for export shares in our baseline model in Table 4. Indeed, the magnitudes of exchange-rate change, volatility, and foreign-income effects on the real exports of firms are found to be larger than on the export shares of firms. Thus, measuring exporting behavior using either export share or real exports yields similar average exchange-rate and foreign-income effects across firms.

Table 9 Robustness check I: Using change in real exports as a dependent variable

Variables	Real exports (WPI) I	Real exports (CPI) II
REER change	-0.085*** (0.020)	-0.082*** (0.019)
REER volatility	-0.494*** (0.192)	-0.472*** (0.189)
I* change	0.042*** (0.011)	0.045*** (0.011)
Year dummies	yes	yes
Firm fixed effects	yes	yes
Observations	38175	38175

Note: Robust Standard errors in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Constant is not reported. The dependent variable in Column I is percentage change in total exports deflated by CPI [Consumer Price Index-Industrial Workers (fiscal year 2001=100) in local currency, period average, from the Labor Bureau of India]. In Column II, it is percentage change in total exports deflated by WPI (Wholesale Price Index in local currency, period average).

The two nominal exchange-rate variables are constructed following the same methodology as the REER change and volatility variables. Their estimated effects are presented in Table 10. It is worth noting here that the interventions of the Reserve Bank of India are geared towards managing the rupee's nominal exchange-rate and most media discussions on exchange-rate effects also implicitly refer to nominal exchange-rate policy. Thus, it is possible that firms follow the nominal exchange-rate rather than the real rate.

The estimated coefficient of NEER change is negative and statistically significant across all specifications. With the exception of specification V, the NEER change effects have magnitudes comparable to those in Table 4. The NEER volatility effect, however, is quite different from the REER volatility effect. Specifically, the NEER volatility variable is only statistically significant under specification V and has a positive sign. The positive effect is opposite to the negative effect for REER volatility in Table 4. As noted earlier, the volatility effect on trade is not conclusive from either a theoretical or an empirical point of view. However, it will be of interest in further research to investigate the different real and nominal exchange rate volatility effects on trade.

Table 10 Robustness check II: Effect of NEER change and volatility

Variables	I	II	III	IV	V
NEER change	-0.063*** (0.019)		-0.067*** (0.019)	-0.071*** (0.019)	-0.030* (0.019)
NEER volatility		0.063 (0.073)	-0.034 (0.076)	-0.147 (0.097)	0.216*** (0.075)
I* change				0.025** (0.012)	-0.011 (0.012)
Year dummies	no	no	no	no	yes
Firm fixed effects	yes	yes	yes	yes	yes
Observations	33132	33132	33132	33132	33132

Note: Robust standard errors in parentheses; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$. Constant is not reported. The table presents results of estimating Equation (1). I^* is the level of foreign income (proxied by trade-share-weighted average of incomes of India's principal trading partners) that represents growth in overseas markets.

In addition to these results, we conducted several other analyses to assess the robustness of the real exchange-rate effect. To conserve space, we briefly discuss the results of these additional exercises below. Details are available on request.

We excluded firms with less than three years of data and re-estimated Equation (3) with firm-level control variables and macro economic factors. We found that results for the sample where firms are present for a sufficient time are qualitatively similar to those reported in previous sections.

Next, as an alternative way to assess the implications of the global financial crisis, we re-estimated the canonical trade equation using two non-overlapping samples of 2000 to 2007 and 2008 to 2010. In both sub-samples, the REER change effect was found to be statistically negative. The effect was stronger between 2000 and 2007 (coefficient estimate = -0.15) than during the crisis period (coefficient estimate = -0.04). The strong negative exchange-rate effect in the pre-2008 period was also recorded for other specifications such as those that allow for asymmetric responses reported in Table 7. The small post-2008 estimated exchange-rate effect is possibly due to both the relatively short sample period and the high volatility experienced during the crisis period.

Finally, we assessed the possible roles of several other macro variables that could conceivably affect exporting behavior of firms. Specifically, we considered the effects of a

terms-of-trade variable, a crude oil price measure, and the VIX index that represents overall market volatility. The encouraging finding is that, in the presence of these variables, the real exchange effect maintains its significance and, usually, is stronger than the one reported in Table 4.

7 Concluding remarks

India, one of the BRIC countries, launched a series of globalization and liberalization reform initiatives in 1991. Undeniably, India's economy has become increasingly linked to the rest of the world and now has a considerable presence as an exporter in the global economy. Exports corresponded to 7% of GDP in 1990 and nearly 23% in 2010.

Against this backdrop, we investigate the exchange-rate effects on exports using Indian firm-level data. Specifically, we use detailed data on a sample of Indian non-financial sector firms for the period 2000–2010 to analyze the effects of REER change and volatility on the share of firm output being exported. During the sample period, exports registered a remarkable increase and the REER exhibited a steady appreciation trend (with the exception of the post-crisis period of 2009 when there was a sharp depreciation).

We asked whether such appreciation has an adverse impact on a firm's exports as predicted by traditional economic theory.

Our basic empirical analysis reveals that indeed over the sample period a currency appreciation had a strong and significant negative impact on Indian firms' export shares. REER volatility was also found to have a negative effect on a firm's export decision. In that sense, our results are in line with those reported by other studies using firm-level data from developed countries, as well as Indian studies that use aggregate export data.

While firm-level accounting information has limited implications for the interaction between exchange-rate and exports, labor cost apparently can amplify exchange-rate effects on trade. Further, there is evidence that Indian firms respond asymmetrically to exchange-rate movements. The REER change effect is likely to be driven more by the negative appreciation effect, than by the depreciation effect. Indian firms with smaller export shares were also found to have stronger responses to both REER change and volatility than firms with larger than median export shares. Firms that export services seem to be more affected by the exchange-rate than firms that export goods.

These results have several important policy implications. Given the Indian economy's dependence on exports and the Reserve Bank of India's policy of a managed exchange-rate, our empirical results indicate that currency appreciation and currency volatility, in general, have an adverse effect on Indian exporters. This circumstantial evidence suggests that if the goal of policymakers is to promote exports, especially in light of India's faltering GDP growth at the moment, they should focus their efforts on restraining rupee appreciation and reducing volatility. Indeed, it appears that the Reserve Bank of India has been pursuing such an asymmetric policy of intervening to prevent appreciation over the last decade or so (Sen Gupta and Sengupta, 2013).

As noted earlier, exchange-rate policy seems to have a stronger impact on some firms than others. Our exercise here indicates, for example, that the effect tends to be stronger for firms with smaller export shares as well as for firms engaged in export of services. Thus, a simple policy of managing the rate of appreciation to promote exports may have the unintended consequences by creating imbalances between different types of firms.

However, an exchange-rate policy that simply strives to prevent appreciation may fall victim to the so-called Penn effect, which briefly stated, suggests that appreciation of the currency goes hand in hand with economic growth. During the sample period, India's average per capita real economic growth was 5.54%. If we assume the income effect on real exchange-rate is 0.25 (Cheung, Chinn, and Fujii, 2007), then the implied rate of appreciation is $1.38\% = 0.25 * 5.54\%$. Since the actual average real exchange-rate appreciation of 1.35% recorded during the sample is smaller than the implied appreciation, one could argue that the observed exchange-rate is undervalued. Thus, a policy of preventing appreciation may be counter-productive. Indeed, it is worth noting in this context that the 2001–2002 crisis in Argentina (also an emerging economy) was arguably the result of attempts by its policymakers in the 1990s to rein in currency appreciation during a period of high economic growth.

The numerical example discussed above is mainly for illustration. We know that the income effect on real exchange-rate is estimated with considerable amount of uncertainty. The important point to note here is that, for a fast-growing country, the trade effect of the exchange-rate is not easy to pin down. A discussion that focuses exclusively on exchange-rates runs the risk of overlooking other factors that may hinder India's exports. For instance, our results are suggestive of some kind of cross-market interaction such as higher labor costs potentially magnifying negative exchange-rate effects. Thus, policymakers

should not focus exclusively on an exchange-rate policy to promote India's export activity. India's export performance can likely benefit from a host of catalytic factors such as investments in infrastructure including highways and ports, liberalization of the labor market, and a concerted policy to promote manufacturing industries.

On the other hand, the negative REER volatility effect consistently found in most of our reported regressions deserves closer scrutiny. Reducing REER volatility seems likely to enhance export performance. In addition to improved hedging instruments, policies could be devised to alleviate volatility. As noted by Aizenman, Edwards, and Riera-Crichton (2012), the exchange rate volatility reducing policies usually come with costs including those associated with accumulating a high level of international reserves.

To summarize, a finer classification of firms and exchange-rate movements suggests that the exchange-rate effects on exports of firms are more complex than simple textbook prescriptions. To gain further insight into the exporting behavior of Indian firms, and especially to draw applicable policy lessons from an empirical analysis, future research may be warranted to examine factors underlying the asymmetric responses of exporting firms to exchange-rate movements. Access to destination-specific and sector-specific firm-level export and price data, extension of the sample period and application of alternative estimation techniques, may provide opportunities to further enhance our understanding of the exporting behavior of Indian firms.

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Data appendix

Macro Variables	Definitions/Descriptions	Data Sources
REER	Real-effective-exchange-rate-weighted averages of bilateral exchange rates adjusted by relative consumer prices. Weighting pattern is time varying, and weights are based on trade for 2008–2010. 2010 is the index base year.	Bank of International Settlements (BIS) database
REER change	Annual percentage change in REER indices.	BIS
REER volatility	Annual standard deviation of monthly REER indices.	BIS
Change in nominal wages	Annual percentage change in hourly wages in local currency, over previous year.	Yearbook of Labor statistics from the International Labor Organization (ILO)
Change in real wages	Annual percentage change in hourly wages in local currency adjusted for inflation, over previous year.	Derived from ILO labor statistics
Real per capita GDP growth rate	Annual percentage change in real per capita GDP.	World Development Indicators (WDI) database.
Change in foreign income level	Change in trade-share-weighted incomes of top ten trading partners.	GDP data of trade partners from World Development Indicators (WDI). Trade share data from the Export-Import data bank maintained by India's Department of Commerce in the Ministry of Commerce and Industry
Change in share of world exports	Percentage change in ratio of world exports to world GDP.	World Development Indicators (WDI) database.

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