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Equilibrium exchange rates in Central and Eastern Europe: A meta-regression analysis

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Balázs Égert – László Halpern

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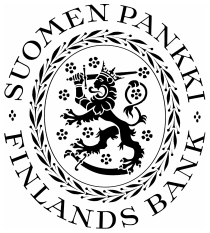

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Equilibrium exchange rates in Central and Eastern Europe: A meta-regression analysis

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Contents

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Balázs Égert – László Halpern

Equilibrium exchange rates in Central and Eastern Europe: A meta-regression analysis

Abstract

This paper analyses the ever-growing literature on equilibrium exchange rates in the new EU member states of Central and Eastern Europe in a quantitative manner using meta-regression analysis. The results indicate that the real misalignments reported in the literature are systematically influenced, inter alia, by the underlying theoretical concepts (Balassa-Samuelson effect, Behavioural Equilibrium Exchange Rate, Fundamental Equilibrium Exchange Rate) and by the econometric estimation methods. The important implication of these findings is that a systematic analysis is needed in terms of both alternative economic and econometric specifications to assess equilibrium exchange rates.

JEL: C15, E31, F31, O11, P17.
Keywords: equilibrium exchange rate, Balassa-Samuelson effect, meta-analysis

Balázs Égert – László Halpern

Equilibrium exchange rates in Central and Eastern Europe: A meta-regression analysis

Tiivistelmä

Tässä tutkimuksessa analysoidaan metaregressioiden avulla Keski- ja Itä-Euroopan tasapainovaluuttakursseja käsittelevää kirjallisuutta. Raportoituihin poikkeamiin tasapainovaluuttakursseista vaikuttavat tulosten mukaan systemaattisesti valittu teoreettinen viitekehys ja käytetyt ekonometriset menetelmät. On siis tärkeää analysoida systemaattisesti käytettyjä menetelmiä ja teorioita tasapainovaluuttakursseja koskevia tuloksia tulkittaessa.

Asiasanat: tasapainovaluuttakurssi, Balassa–Samuelson-efekti, meta-analyysi

1 Introduction

Equilibrium exchange rates have consistently drawn the attention of both academic researchers and policy-makers in industrialised countries for the last decade (Cf. Williamson, 1994; MacDonald, 1995, 2000; Stein, 1995, 2002; and Driver and Westaway, 2004). This is all the more true for the new EU member states of Central and Eastern Europe, which started their transformation process from plan to market in the late 1980s and early 1990s (e.g. Halpern and Wyplosz, 1997; and Krajnyák and Zettelmeyer, 1998, for the early 1990s). A straightforward way to analyse the increasing literature on equilibrium exchange rates in Central and Eastern Europe is to have recourse to conventional literature surveys (Égert, Halpern and MacDonald, 2004). However, traditional literature surveys may contain some degree of subjectivity, as pointed out in Stanley (2001) and Florax, de Groot and de Mooij (2002). By contrast, a meta-regression analysis of the literature may provide a less narrative and more statistical interpretation of the existing body of the literature in that it pins down structural characteristics and methodological features of the studies, which cause the estimation results of the individual papers to differ. Although meta-regression analysis has long been used quite extensively in some areas of economics, only few studies concentrate on macroeconomic issues.1 Even fewer studies deal with transition economics2 and especially exchange rate economics. For the latter area, our paper is, to our knowledge, the first meta-regression study.

Applying the meta-regression approach to the eight new EU member states of Central and Eastern Europe[3], we ask a set of questions that are highly relevant for both academic research and policy-making. The questions that we raise are related to the size and the sign of the estimated real misalignments. In particular, the issues to be answered are whether the estimated real misalignment figures depend on the theoretical background and whether the use of time series, cross-sectional, or panel data, and applying different econometric estimation techniques can systematically influence the estimation results.

[1] Stanley (1998) studies the Ricardian equivalence. Knell and Stix (2003, 2004) analyse the robustness of money demand function estimates. Rose (2004) applies meta-analysis to study the trade creation effect of monetary unions.
[2] Djankov and Murrell (2002) analyse enterprise restructuring in transition economies in a quantitative way. Nonetheless, they do not perform proper meta-regression analysis. Fidrmuc and Korhonen (2004) perform meta-regression to analyse the literature regarding the business cycle correlation between countries in Central and Eastern Europe and the euro area.
[3] The Czech Republic, Estonia, Hungary, Latvia, Lithuania, Poland, Slovakia and Slovenia.

The remainder of the paper is structured in the following way. Section 2 describes the concept of meta-regression analysis. Section 3 presents the results and Section 4 concludes.

2 The concept of meta-regression analysis

Meta-analysis has a long tradition in medicine, especially in clinical medical trials, where it is used because medical experiments are costly, usually take long time and are typically conducted on small groups of individuals. The results of such individual trials all over the world can then be pooled together and analysed as a whole using statistical methods. Stanley (2001) cites the example of streptokinase, for which independent trials provided no conclusive evidence on whether it diminishes the risk of heart attack. Nonetheless, several meta-analyses came to the conclusion that it does have a beneficial effect on the heart.

Meta-analysis helped researcher to clarify controversial issues not only in medicine but also in economics, where it has gained more popularity since the 1980s (see e.g. Stanley and Jarrell, 1989, for an early overview of meta-analysis). Labour economics, industrial organisation, health economics and transportation economics are typical examples of areas where meta-analysis has been used extensively since the late 1980s.

According to Weichselbaumer and Winter-Ebmer (2003), "a meta-study (…) allows a quantitative assessment of the literature in a way an econometrician would write a survey". It allows to formulate and subsequently to test hypotheses related to, for example, the size or the sign of a given coefficient estimate. Stanley (2001) claims that "meta-regression analysis can (…) offer specific reasons, based on the studies themselves, why the evidence on a certain question may appear contradictory or overly varied. Such studies can also suggest potentially fruitful lines for future inquiry (…)."

Meta-regression analysis, a type of meta-analysis, typically involves three stages: First, collect all relevant studies. Second, identify the dependent and independent variables and code them. The study-to-study variation of the dependent variable is to be explained by the independent variables, which are structural characteristics and methodological features of the individual studies. The dependent variable contains usually a summary measure, such as the size of the real misalignment in our case or a coefficient estimate, whereas the independent variables are typically dummy variables. Third, regress the dependent

variable on the set of independent variables. Stanley (2001) puts forward that "meta-regression analysis can identify the extent to which the particular choice of methods, design and data affect reported results."

3 The meta-regression analysis

3.1 Setting up the experiments

The studies

As suggested above, the first two steps of a meta-regression analysis are the identification of the relevant papers and the appropriate coding of the variables. Our dataset includes 32 papers, mostly drawn from Égert, Halpern and MacDonald (2004) and completed with several other studies that became available by early 2004. Only papers which analyse the eight new EU member states of Central and Eastern Europe, namely the Czech Republic, Estonia, Hungary, Latvia, Lithuania, Poland, Slovakia and Slovenia, and which investigate the macroeconomic definition of the real exchange rate are considered here. The real exchange rate is defined as foreign relative to domestic price levels ($Q = E \cdot P^* / P$), where E is the nominal exchange rate expressed as units of domestic currencies in one unit of the foreign currency (a decrease/increase is an appreciation/depreciation) and P and P* are the domestic and foreign price levels.

Florax, de Groot and de Mooij (2002) point out that a common problem with studies using meta-analysis is the construction of a representative sample of the literature. Our paper is not confronted with this problem:_ we use the whole sample of papers from the mid-1990s to early 2004 rather than a representative sample of the literature. Appendix Table A1 lists the papers used in this study with their main features.

The dependent variable

The dependent variable is the size of the real misalignment, i.e. the difference between the estimated equilibrium exchange rate and the observed real exchange rate. If a misalignment range is given in a study, the mean of the band is taken as the size of the real misalignment. The surveyed 32 studies provide us with a total of 170 observations for real misalignments from 1990 to 2002. If a paper provides more than one observation, i.e. ob-

servations for several countries, or an observation for a given country derived on the basis of different methods, then all these observations are collected. Stanley and Jarrell (1998) use only one observation per study. Weichselbaumer and Winter-Ebmer (2003) argue that this may involve a large degree of discretion and advocate including all observations available in a given study.

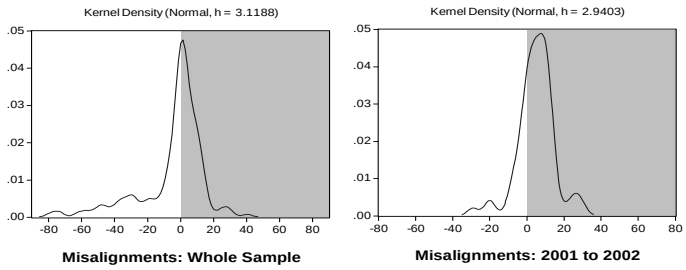
Table 1 below indicates that more than half of the observations, i.e. 88 observations, are from 2001 and 2002. It is reasonable to think that real misalignments obtained for two consecutive years are more comparable than those for the whole sample. For this reason, a sub-sample comprising data only for 2001 and 2002 is employed for the estimations beside the whole sample including misalignments from 1990 to 2002. Table 1 also shows the relative country coverage of our dataset. Of the eight countries analysed, the Czech Republic, Hungary, Poland, Slovenia, Estonia and Latvia are fairly evenly represented both in the whole sample and in the sub-sample. By contrast, Lithuania and Slovakia appear to be somewhat underrepresented. The reported real misalignments range from –79% (undervaluation) to 40% (overvaluation) for the whole sample and from –29% to 30% for the sub-sample. Most of the reported real misalignments are overvaluations (45% for the whole sample and 64% for the sub-sample), as depicted in Figure 1. The large negative figures for the full sample reflect the estimated initial undervaluations. The country-specific kernel density estimates reveal a great deal of heterogeneity across the countries. For the more recent period of 2001 to 2002, most of the countries had overvalued currencies, either in effective terms or vis-à-vis the euro area (or a benchmark proxy), perhaps with the exception of Latvia. Nevertheless, not only the size of the maximum overvaluation but also the mean and the shape of the kernel density estimations vary across the six countries under study.

Table 1. Summary statistics of real misalignments

	1990_2002	2001_2002		1990_2002	2001_2002
N. of Obs	170	88	CZ	15.88%	18.18%
MEAN	-4.64	4.16	HU	18.24%	19.32%
MEDIAN	0.00	3.75	PL	16.47%	18.18%
MAX	40.70	30.00	SK	8.82%	2.27%
MIN	-79.00	-29.00	SI	10.59%	10.23%
STD. DEV.	19.86	9.21	EE	14.71%	14.77%
OVERVALUED	44.71%	63.64%	LV	10.00%	12.50%
UNDERVALUED	35.29%	18.18%	LT	5.29%	4.55%
FAIRLY VALUED	20.00%	18.18%	Total	**100.00%**	**100.00%**
TOTAL	**100.00%**	**100.00%**			

Note: 1990_2002 refers to the whole sample and 2001_2002 stands for the sub-sample. MIS1990 to MIS2002 show the share of the respective years in the sample. CZ, HU, PL, SK, SI, EE, LV and LT denote the Czech Republic, Hungary, Poland, Slovakia, Slovenia, Estonia, Latvia and Lithuania, respectively.

Figure 1. Distribution of real misalignments



Explanatory variables

The explanatory variables are dummy variables, i.e. they take the value of either 0 or 1. An important group of explanatory variables concerns the theoretical background of the studies. The theoretical backgrounds employed are the Balassa-Samuelson effect (BS), the Behavioural Equilibrium Exchange Rate (BEER), the Permanent Equilibrium Exchange Rate (PEER), the Fundamental Equilibrium Exchange Rate (FEER), and the Fundamental Real Exchange Rate (FRER), i.e. the foreign debt-augmented variant of FEER. Although its theoretical background is the risk premium-augmented uncovered interest parity, the BEER approach can be thought of as a statistical approach which aims to link the real exchange rate to a set of economic fundamentals. The equilibrium exchange rate is obtained by plugging in the long-term values of the fundamentals into the estimated equations, and the real misalignment is obtained as the difference between the observed real exchange rate and the estimated "equilibrium" real exchange rate. PEER differs from BEER in that it decomposes directly the estimated long-term cointegration relationship into transitory and

permanent components, with the latter constituting the measure of the equilibrium exchange rate. FEER is a macro model-based approach, in which the equilibrium exchange rate is given by the real exchange rate, which causes the current account to move to its long-term sustainable target, conditioned on the simultaneous attainment of the internal balance usually defined in terms of the nonaccelerating inflation rate of unemployment (NAIRU). FRER differs from FEER in that it also stipulates a foreign debt target to be obtained in the long run.

Given that only two observations are at hand for the NATREX model and the Macroeconomic Balance approach, we decided to ignore them. The single-equation estimate for the NATREX reported in Karádi (2003) is classified as BEER.

As shown in Table 2, BEER is the most commonly used approach with a share of over 50% in the sample. The share of the other approaches differs across the two samples. In the full sample, the BS has a share of 28%, followed by PEER and FEER with about 10% each. In the sub-sample, the second most frequently used approaches are PEER and FEER with a share of about 20% each, while BS ranks at the end with 6%.

Table 2. Theoretical background for real misalignments

	1990_2002	2001_2002
BS	28.24%	5.68%
BEER	52.35%	56.82%
PEER	10.59%	20.45%
MACROMODEL	8.82%	17.05%
Of which		
FEER	3.53%	6.82%
FRER	5.29%	10.23%

In addition to the aforesaid dummies, a score of other, more general variables are also introduced and applied to all specifications. First, a group of explanatory variables are used to capture differences of the econometric estimation techniques (different time series and panel techniques). Second, a class of dummy variables is employed to analyse whether the use of time series, in-sample and out-of-sample estimates and cross-sectional data do matter. A third group of control variables is concerned with the construction of the real exchange rate, i.e. whether it is based on the real effective exchange rate, the real exchange rate vis-à-vis the euro area (or a proxy like Germany or Austria) or the USA, or whether it is based on the CPI, the PPI or real dollar wages. Also, a set of dummy variables is used to control for publication bias, that is, whether published papers produce systematically

higher or lower estimates than those obtained in unpublished papers. To capture year-specific and country-specific misalignments, time and country dummies are used. A detailed definition of the variables is provided in Appendix Table A2.

3.2 Estimation results

Real misalignments

There are two important issues we seek to investigate here. The first one relates to whether the underlying theoretical approaches, i.e. the Balassa-Samuelson effect, BEER, PEER, FEER and FRER have a systematic impact on the size of the real misalignment. Égert, Halpern and MacDonald (2004) provide the time hierarchy of the different theories and argue that although the different approaches are connected with each other, they apply at different time horizons. The estimation results reported in Table 4 lend support to this hypothesis. Comparing the different approaches with the BEER approach[4], it turns out that FEER is significantly different for the whole sample. Regarding the sub-sample, in addition to FEER, PEER also becomes significantly different from BEER. It should be noted that these results are based on the adjusted samples.[5] For the unadjusted data sample, the FRER approach appears to provided different results. A reason why FRER becomes insignificant in the adjusted sample is that its higher values fall in the trimmed upper or lower three percentiles. In general, FEER, FRER and PEER yield higher misalignment figures than BEER.

The second issue we examine is the time hierarchy of the real misalignment estimates. BEER and PEER estimations rest on a single equation which connects the real exchange rate and the fundamentals. Such a specification can be estimated using (a) time series, (b) panel data, and (c) cross-sectional data. If there is a long-term cointegration relationship between the real exchange rate and the fundamentals, real misalignments derived from (a) time series estimates should show a short- and medium-term deviation from the long-term relationship. When using (b) panel data, the estimated deviation of the equilibrium exchange rate from the observed exchange rate may be larger because panel data may

[4] It is always convenient to code the alternative approaches relative to the one with most of the observations. BEER has a relative share of about 50%.

be construed as referring to longer time horizons. The use of in-sample panel data implies that the estimated coefficients reflect some kind of average for a group of transition economies. Thus, the computed real misalignment should be viewed as a medium- to long-term deviation. Out-of-sample data[6] may include either a group of developed countries (e.g. Kim and Korhonen, 2002; Maeso-Fernandez, Osbat and Schnatz, 2004; and Égert, Lahrèche-Révil and Lommatzsch, 2004) or possibly all (market) economies in the world (e.g. Halpern and Wyplosz,1997; and Krajnyák and Zettelmeyer, 1998). Using out-of-sample data including developing countries implies that the equilibrium exchange rate of transition economies behaves like that in developed countries (with which transition economies are making an effort to catch up in the long term), whereas employing out-of-sample data composed exclusively of industrialised economies rests on the assumption that all market economies behave similarly in the (very) long run, as do equilibrium exchange rates. Either way, real misalignments derived from out-of-sample estimates reflect (very) long-run misalignments. Finally, cross-section estimates usually relate the real exchange rate to the dual productivity differential. In such a setting, all variables are expressed in levels rather than the indices commonly used in other BEER estimations. An exception is the paper by Maeso-Fernandez, Osbat and Schnatz (2004), who use level data in a panel setting. Such a bivariate setting is capable of answering the question of how far the real exchange rate is situated from the real exchange rate that would be given by relative productivity levels. Thus, misalignments obtained on the basis of cross-sectional data can be viewed as medium- to long-term misalignments.

Table 3 shows that for the whole sample, the unadjusted data indicate that real misalignments derived on the basis of cross-section and out-of-sample estimations result in higher misalignments than time series estimations. When adjusting for outliers, in-sample estimations appear to yield significantly lower real misalignments than estimations based on time series. The results obtained for the sub-sample 2001 and 2002 should be taken with a grain of salt, because the share of cross-sectional and panel observations is rather limited there. Yet we can find some evidence that in-sample panel estimations provide significantly lower real misalignments than time series estimations. Table 3 also indicates that the

[5] When investigating the determinants of the real misalignments, two equations are estimated. The first one is based on the full sample, whereas the second one is adjusted for possible outliers by trimming the upper and lower three percentiles of the sample. It should also be noted that year-specific and country-specific dummies are always included in the estimated equations.
[6] The expression out-of-sample refers to the fact that the CEECs for which real misalignments are computed are not included in the dataset.

country dummies enter the estimated equations significantly. Given that Hungary is taken as a benchmark, these results suggest that real misalignments in Hungary are systematically different from those of the other countries. More specifically, the real misalignment figures reported in the literature are systematically higher in Hungary than in all other countries except Poland. These results imply, perhaps unsurprisingly, that estimated real misalignments vary across countries.

Table 3. Estimation results for the real misalignments

Dependent variable: Real misalignment								
	1990-2002				2001-2002			
Explanatory	Eq1		Eq2		Eq1		Eq2	
Variables	Full	Adj	Full	Adj	Full	Adj	Full	Adj
C	10.741***	4.836**	10.141***	6.342***	7.783***	4.584***	5.307**	5.404***
Theoretical background								
BS	2.983	1.941			2.812	0.026		
PEER	1.038	0.692			2.213	3.981***	-1.707	-1.305
FEER	7.891	10.203**	12.361*	10.089**	9.794	7.643*	7.428**	0.776
FRER	5.327	1.130	4.048	0.420	7.136**	3.210		
Time series and cross-sectional dimension								
CROSS			14.125**	0.962			1.772	0.472
INSMPL			-0.025	-9.632**			3.297	-5.925*
OUTSMPL			19.118**	-0.527				
Construction of the real exchange rate								
REER	-2.262	-0.600	-0.292	-1.293	-2.204	-2.041		
RER_USD	-2.030	-0.204	-11.019*	-0.433	-5.686	-0.155		
RER_PPI	-8.336*	-7.719**	-7.017	-8.322**	-8.784**	-4.035*		
RER_W	-20.030***	-10.164*	-18.463***	-8.253				
Publication bias								
PUBLI_NAT	-15.554***	-11.934***	-19.382***	-11.736***	-12.596	-7.020		
PUBLI_INT	-7.270	3.127	-1.738	4.464	1.353	4.952		
PUBLI_NO	-5.338	-3.576	-5.965*	-3.877	-5.373**	-0.185		
Country dummies								
CZ	-5.885*	-3.578	-6.005**	-4.210*	0.197	1.641	-0.279	0.763
PL	1.650	2.808	1.687	2.844	1.516	4.144**	1.382	3.540**
SK	-7.050*	-4.956*	-6.748*	-5.102*	10.708*	4.806	10.828	6.041
SI	-3.021	-0.278	-3.348	-0.241	-6.423**	-2.223	-5.452	-2.324
EE	-5.738*	0.093	-5.745*	-0.083	-1.018	-0.383	-1.911	-1.844
LV	-9.938**	-2.594	-8.895**	-2.739	-4.559	-0.388	-10.159***	-4.754**
LT	-12.483**	-3.800	-11.628**	-3.661	-8.756*	-5.688	-10.466**	-6.695*
No: Obs	170	139	170	139	88	69	88	69
R2	0.716	0.622	0.728	0.638	0.511	0.524	0.322	0.405
R2 Adj	0.662	0.531	0.674	0.545	0.384	0.353	0.213	0.277

Note: *,** and *** indicate that the variable is significant at the 10%, 5% and 1% level, respectively. "Full" refers to the raw sample while "adj" is the sample adjusted for possible outliers by trimming the upper and lower three percentiles. Year-specific and country-specific dummies are always included in the equations.

Econometric estimation methods

Some evidence can be found that the econometric estimation methods can influence the size of the derived real misalignment (see Table 4). The Engle-Granger method is used as a common denominator in all equations. There is at least one alternative econometric technique for all reported equations that produces significantly different coefficient estimates. For the full sample, it appears that the pooled and the random effect panel OLS estimators cause systematically different estimates. The result for random effect panel OLS estimators should be treated with caution. The econometric methods are not tested for jointly with other characteristics of the studies because the econometric characteristics sometimes overlap with other characteristics. Most importantly, large panels for the early and mid-1990s are usually estimated using random effect OLS. Hence, the large initial undervaluation detected in these studies may also show up here as a result of different econometric techniques. For the sub-period of 2001–2002, these studies are not included and there are no overlaps between different characteristics. Table 4 reports that it is mostly the Johansen cointegration technique that yields statistically different misalignment estimations. Note that the results do not change if the Johansen technique is taken as a common denominator for the estimations.

Table 4. Econometric techniques

	Dependent variable			
	Real misalignment			
	Eq 1		Eq 2	
	1990-2002		2001-2002	
Expl. Var.	**Full**	**Adj**	**Full**	**Adj**
C	4.262	0.539	0.251	1.471
FMOLS				
DOLS	3.890	7.158	5.230	3.807
ARDL	3.404	5.749	4.744	3.320
JOHANSEN	3.106	2.191	4.862*	4.60***
POLS	-10.711	-12.882**	3.049	-0.912
FE_OLS	-8.077	-7.737		
RE_OLS	-45.393***	-27.378***		
GLS	-7.629	-10.400		
PFMOLS				
PDOLS	8.182	-2.945	7.646	-1.521
PMGE	0.859	-4.947		
MGE				
No. Obs	155	121	73	54
R2	0.708	0.475	0.320	0.441
R2 Adj	0.648	0.330	0.170	0.260

Note: as for Table 4.

4 Conclusion

Using meta-regression analysis, we found important structural differences for the estimated real misalignment obtained for the eight new EU member states from Central and Eastern Europe. We showed that the underlying theoretical background mattered for real misalignment estimates. BEER, PEER and FEER estimates are found to yield significantly different real misalignment estimates. Also, it turned out that the use of time series and in-sample and out-of-sample panels may cause the size of an over- or undervaluation to differ. These findings may be due to the fact that these approaches apply at different time horizons.

Our results have important implications. If one seeks to assess the equilibrium exchange rate of any given economy, a systematic analysis using alternative economic and econometric specifications must be performed because different approaches and techniques turn out to yield systematically different results. In addition, when interpreting the range of the derived real misalignments, the connection between the alternative theoretical and empirical approaches should be carefully analysed.

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Appendix

Table A1. Overview of the studies analysed

	Countries	Approach	Technique
Alberola (2003)	CZ, HU, PL	BEER/PEER	Time series
Alonso-Gamo et al. (2002)	LT	BEER/PEER	Time series
Avallone and Lahrèche (1999)	HU	BEER	Time series
Begg et al. (1999)	CEEC5, EE	BEER	Panel
Beguna (2002)	LV	BEER	Time series
Bitans (2002)	LV	BEER	Time series
Bitans and Tillers (2003)	LV	BEER	Time series
Braumann (1998)	SK	BEER	Time series
Bulir and Smidkova (2004)	CZ, HU, PL, SI	FEER/FRER	--
Burgess et al. (2003)	B3	BEER/PEER	Time series
Cihak and Holub (2001)	CEEC5	BS	Cross-section
Cihak and Holub (2003)	CEEC5, EE	BS	Cross-section
Coudert (1999)	HU	BEER	Panel
Coudert and Couharde (2002)	CEEC5, B3	BS, FEER	Cross-section; ---
Csajbók and Kovács (2003)	HU	FEER	---
DeBroeck and Sløk (2001)	CEEC5, B3	BS	Cross-section
Égert and Lahrèche-Révil (2003)	CEEC5	BEER	Time series
Égert and Lommatzsch (2003)	CEEC5	BEER	Times series, panel
Filipozzi (2000)	EE	BEER	Time series
Halpern-Wyplosz (1997)	CEEC5	BEER	Time series
Hinnosar et al (2003)	EE	BEER	Time series
Karádi (2003)	HU	BEER/NATREX	Time series
Kazaks (2000)	LV	BEER	Time series
Kim and Korhonen (2002)	CEEC5	BEER	Panel
Krajnyák and Zettelmeyer (1998)	CZ, HU, PL, SK, B3	BS, BEER	Cross-section, panel
Lommatzsch and Tober (2002)	CZ, HU, PL	BEER	Time series
Rahn (2003)	CZ, HU, PL, SI, EE	BEER/PEER	Time series
Randveer and Rell (2002)	EE	BS, BEER	Cross-section, time series
Rawdanowich (2003)	PL	BEER	Time series
Smidkova et al. (2002)	CZ, HU, PL, SI, EE	FEER/FRER	--
Vetlov (2002)	LT	BEER	Time series
Vonnák and Kiss (2003)	HU	BEER	Time series/Panel

Note: BS, BEER, PEER, FEER, NATREX denote the theoretical approaches used in the papers. CEEC5 includes the Czech Republic, Hungary, Poland, Slovakia and Slovenia. B3 is the three Baltic states, i.e. Estonia, Latvia and Lithuania. CZ, HU, PL, SK, SI, EE, LV and LT stand for the Czech Republic, Hungary, Poland, Slovakia, Slovenia, Estonia, Latvia and Lithuania, respectively.

Table A2. Codes of the dependent and explanatory variables

	DEPENDENT VARIABLE
MISALIGNMENT	= the point estimate of the real misalignment
	EXPLANATORY VARIABLES
Theoretical background	
BS	=1 if a study uses the Balassa-Samuelson framework
BEER	=1 if a study draws on the Behavourial Equilibrium Exchange Rate approach
MACROMODEL	=1 if a study uses a macromodel
FEER	=1 if a study draws in the Fundamental Equilibrium Exchange Rate approach
FRER	=1 if a study draws on the Fundamental Real Equilibrium Exchange Rate
Estimation methods	
OLS_CR	=1 if a study uses OLS for cross sectional data
EG	=1 if a study uses the Engle-Granger method
FMOLS	=1 if a study uses fully modified OLS
DOLS	=1 if a study uses Dynamic OLS
ARDL	=1 if a study uses Autoregressive Distributed Lags
JOHANSEN	=1 if a study uses the Maximum Likelihood estimator of Johansen
POLS	=1 if a study uses pooled OLS
FE_OLS	=1 if a study uses fixed effect OLS
RE_OLS	=1 if a study uses random effect OLS
GLS	=1 if a study uses generalised least squares
PFMOLS	=1 if a study uses panel fully modified OLS
PDOLS	=1 if a study uses panel dynamic OLS
PMGE	=1 if a study uses the pooled mean group estimator
MGE	=1 if a study uses the mean group estimator
Time series and cross-sectional dimension	
TIMESERIES	=1 if a study uses times series
PANEL	=1 if a study uses panel data
IN_SMPL	=1 if a study uses in-sample panel data
OUT_SMPL	=1 if a study uses out-of-sample panel data
CROSS	=1 if a study uses cross sectional data
Real exchange rates	
REER	=1 if a study uses real effective exchange rate
RER_EURO	=1 if a study uses real exchange rate vis-à-vis a proxy of the euro area
RER_E	=1 if a study uses real exchange rate vis-à-vis the euro area
RER_DE	=1 if a study uses real exchange rate vis-à-vis Germany
RER_AT	=1 if a study uses real exchange rate vis-à-vis Austria
RER_USD	=1 if a study uses real exchange rate vis-à-vis the US
RER_CPI	=1 if a study uses CPI-deflated real exchange rate
RER_PPI	=1 if a study uses PPI-deflated real exchange rate
RER_W	=1 if a study uses dollar wage as a proxy for the real exchange rate
Publication record	
PUBLI	=1 if a study is published in a peer-reviewed journal
PUBLI_INT	=1 if a study is published in an international peer-reviewed journal
PUBLI_NAT	=1 if a study is published in a non-English peer-reviewed journal
PUBLI_WP	=1 if a study appeared as a working paper, is published in a book, conference volume or in a not refereed journal
PUBLI_NO	=1 if a study is a mimeo or conference paper

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