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Article

A Dynamic Programming Approach for Pricing Weather Derivatives under Issuer Default Risk

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Abstract: Weather derivatives are contingent claims with payoff based on a pre-specified weather index. Firms exposed to weather risk can transfer it to financial markets via weather derivatives. We develop a utility-based model for pricing baskets of weather derivatives under default risk on the issuer side in over-the-counter markets. In our model, agents maximise the expected utility of their terminal wealth, while they dynamically rebalance their weather portfolios over a finite investment horizon. Using dynamic programming approach, we obtain semi-closed forms for the equilibrium prices of weather derivatives and for the optimal strategies of the agents. We give an example on how to price rainfall derivatives on selected stations in China in the universe of a financial investor and a weather exposed crop insurer.

Keywords: dynamic programming; pricing; risk management

MSC: 90C39; 91G20; 91B25

1. Introduction

Weather derivatives (WDs) are contingent claims with payoffs determined by future weather events as temperature, snowfall, and rainfall. Hedging with WDs reduces exposure to weather volatility and stabilises profits of a weather exposed agent. Firms operating in energy, tourism, agriculture, and insurance sectors use WDs to hedge their weather risks. Investigations in [Peréz-González and Yun \(2013\)](#) show that weather risk management with WDs leads to an increase in firm value. Moreover, WDs are also attractive for a purely financial investor as their payoffs are acyclic, uncorrelated with financial assets, and, therefore, contribute to portfolio diversification.

Frequently, the structure of agent's weather exposure is complicated, such that, it is more beneficial to purchase a basket of WDs on several underlying weather indices rather than on a single one to manage weather risks in an optimal way. For a financial investor, positions in multiple WDs strengthen the positive portfolio diversification effect. To increase the hedging efficiency and to achieve a higher degree of portfolio diversification, both weather exposed business and financial investors are, therefore, interested in holding a portfolio of multiple WDs. While valuing baskets of such derivatives, market participants should also account for the dependence in the underlying weather indices.

We develop a simple utility-based model for pricing such baskets of customised WDs on multiple dependent underlying indices. We address here the problem of WD pricing in over-the-counter (OTC)

markets, where fewer interested agents transfer their specific non-financial weather risks among themselves through creating and trading customised weather dependent claims. We perceive a model which is flexible enough to handle a wide range of customised WDs and allows dynamic trading in the sense that the agents rebalance their portfolios (or renegotiate the contracts) at some predetermined discrete time points in their finite investment horizon. As in [Çanakoğlu and Özekici \(2009\)](#) we use a dynamic programming approach for portfolio optimisation combined in our case with zero-net-supply condition for WDs to determine their equilibrium prices.

We choose an equilibrium pricing approach based on utility maximisation for several reasons. First, we anticipate that the OTC transactions of WDs involve a limited number of participants. Thus, the interaction of their individual supply and demand functions will give the equilibrium price. Further, due to the market incompleteness standard arguments imposing an existence of a unique pricing measure fail, see [Heath et al. \(1992\)](#). Models based on pure weather dynamics impose restrictive assumptions in order to choose the appropriate equivalent martingale measure, see [Alaton et al. \(2002\)](#), [Benth et al. \(2007\)](#), and [López Cabrera et al. \(2013\)](#). In our model, involved agents construct WD securities in zero net supply, thus, they complete the WD market. Moreover, the chosen utility-based pricing approach naturally reflects the risk attitudes of each agent and facilitates via dynamic programming the derivation of explicit prices for WDs in a multi-period framework.

Most previous works on utility-based WD pricing either fail to price derivatives on multiple dependent underlying indices simultaneously ([Carmona and Diko 2005](#); [Lee and Oren 2010](#); [Leobacher and Ngare 2011](#)) or they impose restrictive assumptions on the dynamics of the underlying weather indices ([Horst and Müller 2007](#); [Chaumont et al. 2006](#)). Consumption based model of [Cao and Wei \(2004\)](#) prices multiple WDs in an extended pure exchange economy of [Lucas \(1978\)](#) on the macroeconomic level and is not feasible for a limited number of weather market participants and a short contract duration typical for OTC markets. We contribute to this literature by providing a model for pricing baskets of customised WDs on multiple dependent underlying indices. In our framework, market participants account for the possible dependence between the weather indexes that determine the payoffs of their weather portfolios. The applicability of our pricing model is not restricted to any particular type of WDs. Various kinds of WDs can be priced in our framework.

Our further contribution is the introduction of counterparty default risk into the pricing model for WDs. The importance of counterparty default risk for pricing was addressed, for example, in [Hull and White \(1995\)](#), [Jarrow and Turnbull \(1995\)](#), and [Wu and Chung \(2010\)](#). We show that the introduction of a non-zero issuer default probability significantly depresses the demand for WDs through correcting expectations of future portfolio payoffs in utility terms downwards.

We illustrate our approach on pricing rainfall derivatives. These derivatives can be used to hedge agricultural volumetric risks, see [Musshoff et al. \(2010\)](#). Traditionally, volumetric risk hedging in agriculture is taken over by the crop insurance as [Glauber et al. \(2002\)](#). Therefore, a crop insurer frequently faces indirect losses caused by rainfall surplus/deficit through its impact on crop production. Whenever the weather exposed income of such an insurer depends on rainfall outcomes in a number of geographical sites where the insured farmers are operating, a basket of rainfall derivatives should be used for hedging the risks.

We give the pricing example for China. Chinese farmers are exposed to pronounced weather risks as [Turvey and Kong \(2010\)](#). According to [The World Bank \(2007\)](#) the existing agricultural insurance schemes are too expensive for Chinese agricultural producers. Trading WDs can play an important role in transferring a part of the weather exposure of crop insurer to financial markets and so make crop insurance affordable for Chinese farmers.

Our empirical analysis addresses the effects of the increasing investment horizon, default risk, market volatility, and capital costs on the demand and on the supply for WDs.

The structure of the paper is as follows. In Section 2 we, first, derive the multi-asset dynamic pricing model for WDs in the absence of any default risk. Next, we extend the model to account for counterparty default risk and a possibility of an alternative financial investment. Section 3 shows an

example of pricing rainfall derivatives using historical rainfall data for China and features a discussion of further applications. In Section 4 we summarise our results.

2. Dynamic Pricing Model for Weather Derivatives

In this section, we construct a dynamic pricing model for WDs with multiple assets. First, we introduce a market design and notation. Then, we derive equilibrium prices for WDs without any default risk. Next, we study the implications of counterparty default risk and of an alternative financial investment.

2.1. Assumptions and Notation

1. *Assets.* There are WDs on S weather indices (at different geographical sites and/or on different weather events) that are priced at times $t = 0, 1, \dots, T - 1$. At T the payoff of each WD is determined and the cash settlement takes place. The non-negative price of the WD on underlying s in time t is denoted as $W_{t,s}$ where $s = 1, \dots, S$ and $t = 1, \dots, T - 1$. The final value $W_{T,s}$ corresponds to the non-negative payoff of the s th WD. We denote the vector of prices at t as $W_t = (W_{t,1}, W_{t,2}, \dots, W_{t,S})^\top$. Besides the WDs, a risk free asset B_t with a constant per period return r is available. Trading with B_t is not restricted in any way, that is, unlimited borrowing and lending at the interest r in each t is allowed. We assume there is no transaction costs on the asset market. No capital addition or withdrawals are possible throughout the investment horizon, such that the agents are exposed to self-financing constraints.
2. *Agents.* There are $J + 1$ heterogeneous market participants, indexed by i , with the risk preferences described by the exponential utility function of the form $U_i(x) = -\exp(-a_i x)$, where $a_i > 0$ is the risk aversion of agent i . All agents have the same multi-period investment horizon of length T . They invest at $t = 0$ and they consume their terminal wealth at $t = T$. At $t = 1, \dots, T - 1$ agents rebalance their weather portfolios and renegotiate the prices for WDs. All agents are endowed with an initial wealth of zero monetary units. We distinguish between J buyers, indicated by subscript j , $j = 1, \dots, J$, who hedge weather exposure of their random income I_j , and a purely financial investor, indicated by subscript m , who issues WDs. Each buyer holds a basket of WDs on the relevant weather indices to hedge weather caused fluctuations in her profits. The issuer holds positions in all S WDs. A portfolio of agent i includes $\alpha_{i,t} = (\alpha_{i,t,1}, \dots, \alpha_{i,t,S})^\top$ shares of the corresponding WDs and $\beta_{i,t}$ shares of the asset B_t . Both $\alpha_{i,t}$ and $\beta_{i,t}$ are real valued, that is, all assets are perfectly divisible and short sales are allowed. We denote the value of i th agent's portfolio at time t as $V_{i,t}$, where $V_{i,t} = \alpha_{i,t}^\top W_t + \beta_{i,t} B_t$. In each period t of the investment horizon, the agents maximise their expected utility of the terminal wealth with the available WDs and attain their demand and supply for the WDs. That is, in each period $t < T$ every agent i determines her self-financing trading strategy $(\alpha_{i,t+1}, \beta_{i,t+1})_{t=0,1,\dots,T}$, in particular, she constructs the optimal hedging portfolio given the state of the system at time t . Partial market clearing with respect to WDs determines the equilibrium prices for the WDs.
3. *State.* The observable state of the system at time t , denoted by \mathcal{W}_t , contains values of the underlying weather indices at t and the variables regarding default risk at t . The random state \mathcal{W}_{t+1} is characterised by the conditional distribution function $\Phi_t(w', w) = \Pr(\mathcal{W}_{t+1} \leq w' | \mathcal{W}_t = w)$. We assume that this transition function Φ satisfies the Feller property [Stokey et al. \(1989\)](#). Expectation taken with respect to $\Phi_t(\cdot)$ is denoted by $E_t(\cdot)$.

Each agent $j \in J$ is faced with the following discrete time stochastic control system:

$$V_{j,t+1} = g_{j,t} \{V_{j,t}, (\alpha_{j,t,s})_{s \in S_j}, \mathcal{W}_{j,t}\}, \quad t = 0, 1, \dots, T \tag{1}$$

where $V_{j,t}$ incorporates portfolio value of agent j with $V_{j,0} = 0$. $(\alpha_{j,t,s})_{s \in S_j}$ are controls of the agent in the system (1). The law of motion $g_{j,t} : \mathbb{R}^{|S_j|+1} \times \mathbb{R}^{|S_j|} \times \mathbb{R}^{|S_j|} \mapsto \mathbb{R}^{|S_j|+1}$ maps to the next state of the stochastic system.

2.2. Pricing WDs without Default Risk

We now derive a dynamic pricing model for WDs with multiple assets under the market design described by assumptions 1 to 3. For the moment, we assume that there is no default risk. The terminal wealth of buyer j at T is her profit $\Pi_{j,T}$:

$$\begin{aligned} \Pi_{j,T} &= I_j + \alpha_{j,T}^\top W_T + \beta_{j,T} B_T \\ &= I_j + V_{j,T} \end{aligned} \tag{2}$$

with I_j being a random income that depends on some weather indices entering the final payoff W_T . $\beta_{j,T} B_T$ and $\alpha_{j,T}^\top W_T$ are the payoffs of the risk free asset and of the basket of the WDs; together they constitute $V_{j,T}$, the terminal portfolio value of buyer j . Let $\alpha_{j,t+1:T}$ denote trading strategies of agent j from $t + 1$ to T . Portfolio choice problem of buyer j in each $t = 0, 1, \dots, T - 1$ is:

$$\begin{aligned} \max_{\alpha_{j,t+1:T} \in \mathbb{R}^{S \times (T-t)}} & \mathbb{E}_t \{U_j(\Pi_{j,T})\} \\ \text{s.t.} & \alpha_{j,t+1}^\top W_t + \beta_{j,t+1} B_t - V_{j,t} = 0. \end{aligned} \tag{3}$$

That is, in each period of the investment horizon T buyer j maximises expected utility of her terminal wealth with respect to all future trading strategies, subject to a self-financing portfolio.

The terminal wealth of investor m at T is:

$$\Pi_{m,T} = -\alpha_{m,T}^\top W_T + \beta_{m,T} B_T = V_{m,T} \tag{4}$$

with $\alpha_{m,T}^\top W_T$ and $\beta_{m,T} B_T$ being the payoffs of the WDs portfolio and the risk free asset respectively. Investor's portfolio choice problem in each $t = 0, 1, \dots, T - 1$ is:

$$\begin{aligned} \max_{\alpha_{m,t+1:T} \in \mathbb{R}^{S \times (T-t)}} & \mathbb{E}_t \{U_m(\Pi_{m,T})\} \\ \text{s.t.} & \alpha_{m,t+1}^\top W_t - \beta_{m,t+1} B_t + V_{m,t} = 0. \end{aligned} \tag{5}$$

In $t < T$ investor m maximises expected utility of her terminal wealth with respect to all future trading strategies, subject to a self-financing portfolio.

Note, that under assumptions 1 to 3, the constraints in the optimization problems (3) and (5) are non-empty, compact and continuous in V_{\cdot} and \mathcal{W}_{\cdot} , $U_{\cdot}(\cdot)$ is continuous and bounded, and transition function Φ is Feller (by assumption). Thus, continuous law of motion g_{\cdot} ensures the existence of optimal solutions to the problems above and continuity of the value function (see [Stokey et al. 1989](#), p. 62).

Following [Pennacchi \(2008\)](#) we solve the multi-period portfolio choice problems (3) and (5) using dynamic programming. Let $J_{i,t}(V_{i,t}, \alpha_{i,t+1:T}, \mathcal{W}_t)$ denote the expected utility of agent i in time t . It depends on the agent's current portfolio value $V_{i,t}$, the current state \mathcal{W}_t , and the trading strategies $\alpha_{i,t+1:T}$ in the time span $t + 1$ to T (here indicated by subscript $t + 1 : T$). Let $J_{i,t}^*(V_{i,t}, \mathcal{W}_t)$ be the expected utility maximised with respect to the strategies of agent i from $t + 1$ on. Formally, in each t $J_{i,t}^*(V_{i,t}, \mathcal{W}_t)$ is obtained by:

$$J_{i,t}^*(V_{i,t}, \mathcal{W}_t) = \max_{\alpha_{i,t+1:T} \in \mathbb{R}^{S \times (T-t)}} \{J_{i,t}(V_{i,t}, \alpha_{i,t+1:T}, \mathcal{W}_t)\}. \tag{6}$$

According to the principle of dynamic programming the expected utility of agent i in t can be rewritten as:

$$J_{i,t}(V_{i,t}, \alpha_{i,t+1}, \mathcal{W}_t) = \mathbb{E}_t [J_{i,t+1}^* \{V_{i,t+1}(\alpha_{i,t+1}), \mathcal{W}_{t+1}\}]. \tag{7}$$

Combining (6) and (7), we obtain a recursive expression for $J_{i,t}^*(V_{i,t}, \mathcal{W}_t)$:

$$J_{i,t}^*(V_{i,t}, \mathcal{W}_t) = \max_{\alpha_{i,t+1} \in \mathbb{R}^S} E_t [J_{i,t+1}^*\{V_{i,t+1}(\alpha_{i,t+1}), \mathcal{W}_{t+1}\}] \tag{8}$$

with the boundary condition at T

$$J_{i,T}^*(V_{i,T}, \mathcal{W}_T) = U_i(I_i + \alpha_{i,T}^\top \mathcal{W}_T + \beta_{i,T} B_T), \tag{9}$$

and with $\beta_{i,T} B_T = V_{i,T} - \alpha_{i,T}^\top \mathcal{W}_T$.

Starting with $t = T - 1$ and using (8), we recursively solve for supply and demand in each t for all the WDs, denoted in vector form as:

$$W_t^i(\alpha_{i,t+1}) = \{W_{t,1}^i(\alpha_{i,t+1}), \dots, W_{t,S}^i(\alpha_{i,t+1})\}^\top, \quad i = j, m. \tag{10}$$

Imposing partial market clearing for WD securities (zero net supply), $\sum_{j=1}^J \alpha_{j,t+1,s}^* = \alpha_{m,t+1,s}^*$, delivers the equilibrium prices $W_t = W_t^m(\alpha_{m,t+1}^*)$ for all $t = 0, \dots, T - 1$, where $*$ identifies the optimal values. Specific forms of buyer’s demand and investor’s supply are given in Propositions 1 and 2 below.

Proposition 1. *Let the utility function of the buyer j be of the exponential form with risk aversion $a_j > 0$. Let the assumptions 1 to 3 hold and let $\Theta_{j,t} < \infty$ in (13) hold for $t \leq T$. Then, j th buyer’s reverse demand for WD s , $W_{t,s}^j(\alpha_{j,t+1})$, and her optimal utility level $J_{j,t}^*(V_{j,t}, \mathcal{W}_t)$ are given recursively by:*

$$W_{t,s}^j(\alpha_{j,t+1}) = \frac{E_t \{ \exp(-a_j \alpha_{j,t+1}^\top W_{t+1} R^{T-(t+1)}) \Theta_{j,t+1} W_{t+1,s} \}}{E_t \{ \exp(-a_j \alpha_{j,t+1}^\top W_{t+1} R^{T-(t+1)}) \Theta_{j,t+1} \} R}, \tag{11}$$

$$J_{j,t}^*(V_{j,t}, \mathcal{W}_t) = - \exp(-a_j V_{j,t} R^{T-t}) \Theta_{j,t}, \tag{12}$$

$$\Theta_{j,t} = \exp(a_j R^{T-t} \alpha_{j,t+1}^{*\top} W_t) E_t \{ \exp(-a_j R^{T-(t+1)} \alpha_{j,t+1}^{*\top} W_{t+1,s}) \Theta_{j,t+1} \}, \tag{13}$$

for $t = 0, \dots, T - 1$, $\Theta_{j,T} = \exp(-a_j I_j)$ and $R = 1 + r$.

Proof. See Appendix A. \square

Proposition 2. *Let the utility function of the investor be of the exponential form with risk aversion $a_m > 0$. Let the assumptions 1 to 3 hold and let $\Theta_{m,t} < \infty$ in (16) hold for $t \leq T$. Then, investor’s reverse supply for WD s , $W_{t,s}^m(\alpha_{m,t+1})$, and her optimal utility level $J_{m,t}^*(V_{m,t}, \mathcal{W}_t)$ are given recursively by:*

$$W_{t,s}^m(\alpha_{m,t+1}) = \frac{E_t \left[\exp \left\{ a_m (\alpha_{m,t+1}^\top W_{t+1,s} R^{T-(t+1)}) \Theta_{m,t} W_{t+1,s} \right\} \right]}{E_t \left\{ \exp \left(a_m \alpha_{m,t+1}^\top W_{t+1} R^{T-(t+1)} \right) \Theta_{m,t} \right\} R}, \tag{14}$$

$$J_{m,t}^*(V_{m,t}, \mathcal{W}_t) = - \exp(-a_m V_{m,t} R^{T-t}) \Theta_{m,t}, \tag{15}$$

$$\Theta_{m,t} = \exp(-a_m \alpha_{m,t+1}^{*\top} W_t R^{T-t}) E_t \left\{ \exp(a_m R^{T-t+1} \alpha_{m,t+1}^{*\top} W_{t+1}) \Theta_{m,t+1} \right\} \tag{16}$$

for $t = 0, \dots, T - 1$, $\Theta_{m,T} = 1$ and $R = 1 + r$.

Proof. The proof is very similar to the proof of Proposition 1 and is omitted. \square

The results above show that the reverse demand and supply for the s th WD at time t are determined by the interaction of the terms proportional to the marginal expected utility of the next period as well as future expected utilities in $\Theta_{i,t+1}$ that embrace optimal future trading behaviour, $\alpha_{i,t+1:T}$. Moreover, the reverse demand and supply are both influenced by capital costs or risk-free rate r . Changes in r will, therefore, influence the resulting WD prices.

The expressions obtained are of a semi-closed form, since the expectations entering (11) and (14) have, in general, no explicit representation and have to be computed numerically using, for example, Monte Carlo techniques.

2.3. Default Risk

As we already mentioned, OTC derivatives are subject to counterparty default risk or credit risk as termed in Golden et al. (2007). In this section we consider incorporation of issuer default risk into the pricing model for WDs.

As in Golden et al. (2007), we consider dichotomous default risk modelled by a payoff proportion parameter λ_t with values in $\{0, 1\}$. While $\lambda_t = 1$ means that the counterparty fully meets its obligations in t , $\lambda_t = 0$ inherits the situation where the counterparty does not perform at all. We assume that λ_t is independent of all other random variables in the model, and $P(\lambda_{t+1} = 0) = p_t$, $P(\lambda_{t+1} = 1) = 1 - p_t$, where p_t is the probability of investor's default at $t + 1$. Parameters λ_t and p_t are known to the buyers by t . Once the investor defaults, she is not going to meet any obligations in subsequent periods, and $\lambda_{t:T} = 0$.

Clearly, under investor's default risk, the payoff vector of the WDs is a multiple of λ_T , that is, $\lambda_T W_T$. If the investor has met all her obligations up to period $t < T$, then the contribution of the next period WD prices is $\lambda_{t+1} W_{t+1}$. Hence, the j th buyer's demand for WD s in (11) has to be modified to account for the default risk in the following way:

$$W_{t,s}^j(\alpha_{j,t+1}) = \frac{E_t\{\exp(-a_j \alpha_{j,t+1}^\top \lambda_{t+1} W_{t+1} R^{T-(t+1)}) \tilde{\Theta}_{j,t+1} \lambda_{t+1} W_{t+1,s}\}}{E_t\{\exp(-a_j \alpha_{j,t+1}^\top \lambda_{t+1} W_{t+1} R^{T-(t+1)}) \tilde{\Theta}_{j,t+1}\} R} \tag{17}$$

with

$$\begin{aligned} \tilde{\Theta}_{j,t} &= \exp(a_j R^{T-t} \alpha_{j,t+1}^{*\top} W_t) \\ &\cdot E_t\{(1 - p_t) \exp(-a_j R^{T-(t+1)} \alpha_{j,t+1}^{*\top} \lambda_{t+1} W_{t+1,s}) \tilde{\Theta}_{j,t+1} + p_t \Theta_{j,T}\}, \end{aligned} \tag{18}$$

and $\Theta_{j,T} = \exp(-a_j I_j)$.

To derive (17) we consider all probable outcomes of λ_{t+1} at each t . Thus, in $T - 1$, if $\lambda_{T-1} = 1$ we solve the maximisation problem (3) with the payoff of the WDs scaled by λ_T and equal to $\lambda_T W_T$. As a result, we find $\alpha_{j,T}^* > 0$, $W_{T-1} > 0$, and $J_{j,T-1}^*(V_{j,T-1}, \mathcal{W}_T)$ which is equal to:

$$J_{j,T-1}^*(V_{j,T-1}, \mathcal{W}_{T-1}) = - \exp\{-a_j V_{j,T-1} R\} \Theta'_{j,T-1} \tag{19}$$

with

$$\Theta'_{j,T-1} = \exp\{a_j \alpha_{j,T}^{*\top} W_{T-1} R\} E_{T-1}[\exp\{-a_j (\alpha_{j,T}^{*\top} \lambda_T W_T)\} \Theta_{j,T}]. \tag{20}$$

If $\lambda_{T-1} = 0$ $\alpha_{j,T}^* = 0$, $W_{T-1} = 0$, and

$$J_{j,T-1}^*(V_{j,T-1}, \mathcal{W}_{T-1}) = - \exp\{-a_j V_{j,T-1} R\} \Theta'_{j,T-1}, \tag{21}$$

where $\Theta'_{j,T-1} = E_{T-1}(\Theta_{j,T})$.

We move one period backwards to $T - 2$. If now $\lambda_{T-2} = 1$ we maximise:

$$\begin{aligned} J_{j,T-2}(V_{j,T-2}, \alpha_{j,T-1}, \mathcal{W}_{T-2}) &= E_{T-2}\{J_{j,T-1}^*(V_{j,T-1}, \mathcal{W}_{T-1})\} \\ &= E_{T-2}[- \exp\{-a_j V_{j,T-1} R\} \tilde{\Theta}_{j,T-1}], \end{aligned} \tag{22}$$

$$\tilde{\Theta}_{j,T-1} = \exp\{a_j \alpha_{j,T}^{*\top} W_{T-1} R\} E_{T-1}[(1 - p_{T-1}) \exp\{-a_j (\alpha_{j,T}^{*\top} \lambda_T W_T)\} \Theta_{j,T} + p_{T-1} \Theta_{j,T}]. \tag{23}$$

Note, that the expectation in the first line of (22) is taken under the joint distribution of λ_{T-1} and the other random variables of the model, and in the second line—only with respect to the joint distribution of the later ones. $J_{j,T-2}^*(V_{j,T-2}, \mathcal{W}_{T-2})$ is in this case:

$$J_{j,T-2}^*(V_{j,T-2}, \mathcal{W}_{T-2}) = -\exp\{-a_j V_{j,T-2} R\} \Theta'_{j,T-2} \tag{24}$$

with

$$\Theta'_{j,T-2} = \exp(a_j \alpha_{j,T-1}^{*\top} W_{T-2} R) E_{T-2} [\exp\{-a_j (\alpha_{j,T-1}^{*\top} \lambda_{T-1} W_{T-1})\} \tilde{\Theta}_{j,T-1}] \tag{25}$$

If $\lambda_{T-2} = 0$ then $\alpha_{j,T-1}^* = 0$, $W_{T-2} = 0$, and

$$J_{j,T-2}^*(V_{j,T-2}, \mathcal{W}_{T-2}) = -\exp\{-a_j V_{j,T-2} R\} \Theta'_{j,T-2}, \tag{26}$$

where $\Theta'_{j,T-2} = E_{T-2}(\Theta_{j,T})$. Then, the following expectation reads:

$$E_{T-3}\{J_{j,T-2}^*(V_{j,T-2}, \mathcal{W}_{T-2})\} = E_{T-3}[-\exp\{-a_j V_{j,T-2} R\} \tilde{\Theta}_{j,T-2}] \tag{27}$$

with

$$\begin{aligned} \tilde{\Theta}_{j,T-2} = & \exp(a_j \alpha_{j,T-1}^{*\top} W_{T-1} R) \\ & \cdot E_{T-2}[(1 - p_{T-2}) \exp\{-a_j (\alpha_{j,T}^{*\top} \lambda_T W_T)\} \tilde{\Theta}_{j,T-1} + p_{T-2} \Theta_{j,T}]. \end{aligned} \tag{28}$$

By backward induction we obtain (17) and (18).

From (17) we observe that a non-zero issuer default probability influences adversely the demand for the WDs through correcting the expectations of the future portfolio values and the marginal expected utility of the WD payoffs downwards.

2.4. Alternative Investment

Frequently, the model assumptions implying that the agents' portfolios contain only the risk free asset besides WDs do not hold. This will often be the case for the investor's portfolio. In the following, we relax this restriction and allow the investor to invest in the financial market.

Let's amend the assumptions made in Section 2.1 with the following:

- 1a *Assets.* Let assumption 1. hold. Let F_t be a quoted price of an exchange traded financial asset at time t . While F_t is given, F_{t+1} is random, bounded, and predictable at t . Trading with F_t is not restricted in any way, that is, short and long positions in the asset in each t are possible. We assume there is no transaction costs on the asset market. As before, no capital addition or withdrawals are possible throughout the investment horizon, such that the agents are exposed to self-financing constraints.

For example, F_t can be a share value of an exchange traded fund tracking some financial portfolio, or it can represent the value of the market portfolio itself at t .

- 2a *Agents.* Let assumption 2. hold. Now, issuer m holds additionally $f_{m,t}$ shares of the exchange traded financial asset with exogenous price F_t . Also, $f_{\cdot, \cdot}$ is real valued, that is, all assets are perfectly divisible and short sales are allowed. The value of the issuer's portfolio at time t becomes $V_{m,t} = \alpha_{m,t}^\top W_t - f_{m,t} F_t + \beta_{m,t} B_t$.

As before, in each period t of the investment horizon, agents maximise their expected utility of the terminal wealth with the available WDs and attain their demand and supply for the WDs. That is, in each period $t < T$ issuer m determines her self-financing trading strategy $(\alpha_{m,t+1}, \beta_{m,t+1}, f_{m,t+1})_{t=0,1,\dots,T}^\top$, in particular, she constructs the optimal hedging portfolio given the state of the system at time t . Partial market clearing with respect to WDs determines the equilibrium prices for the WDs.

3a *State.* Let assumption 3. hold. The observable state of the system at time t , denoted by \mathcal{W}_t , contains additionally the quoted price F_t . The random state \mathcal{W}_{t+1} is characterised by the conditional distribution function Φ_t . Expectation taken with respect to Φ_t is denoted by $E_t(\cdot)$.

Under assumptions 1a-3a, the terminal wealth of investor m at T is:

$$\Pi_{m,T} = -\alpha_{m,T}^\top W_T - f_{m,T} F_T + \beta_{m,T} B_T = V_{m,T} \tag{29}$$

with $\alpha_{m,T}^\top W_T$, $f_{m,T} F_T$, and $\beta_{m,T} B_T$ being the payoffs of the WDs portfolio, the alternative financial investment, and the risk free asset respectively. Investor's portfolio choice problem in each $t = 0, 1, \dots, T - 1$ is:

$$\begin{aligned} \max_{\alpha_{m,t+1:T} \in \mathbb{R}^{S \times (T-t)}} E_t \{ U_m(\Pi_{m,T}) \} \\ \text{s.t. } \alpha_{m,t+1}^\top W_t + f_{m,t+1} F_t - \beta_{m,t+1} B_t + V_{m,t} = 0. \end{aligned} \tag{30}$$

As before, in each $t < T$ investor m maximises expected utility of her terminal wealth with respect to all future trading strategies, subject to a self-financing portfolio. The expected utility of the investor in t , $J_{m,t}(V_{m,t}, \alpha_{m,t+1:T}, f_{m,t+1:T}, \mathcal{W}_t)$, is now also a function of her trading strategies on the financial market, denoted as $f_{m,t+1:T}$.

Keeping in mind these modifications and following the steps of Section 2.2, we have to modify investor's supply for WD s in (14) in the following way:

$$W_{t,s}^m(\alpha_{m,t+1}, f_{m,t+1}) = \frac{E_t \left[\exp \left\{ a_m R^{T-(t+1)} (\alpha_{m,t+1}^\top W_{t+1,s} + f_{m,t+1} F_{t+1}) \Theta_{m,t} W_{t+1,s} \right\} \right]}{E_t \left[\exp \left\{ a_m R^{T-(t+1)} (\alpha_{m,t+1}^\top W_{t+1,s} + f_{m,t+1} F_{t+1}) \Theta_{m,t} \right\} \right] R}, \tag{31}$$

$$J_{m,t}^*(V_{m,t}, \mathcal{W}_t) = -\exp(-a_m V_{m,t} R^{T-t}) \Theta_{m,t}, \tag{32}$$

$$\begin{aligned} \Theta_{m,t} &= \exp \{ -a_m R^{T-t} (\alpha_{m,t+1}^{*\top} W_t + f_{m,t+1}^* F_t) \} \cdot \\ E_t \left[\exp \{ a_m R^{T-t+1} (\alpha_{m,t+1}^{*\top} W_{t+1} + f_{m,t+1}^* F_{t+1}) \Theta_{m,t+1} \} \right] \end{aligned} \tag{33}$$

for $t = 0, \dots, T - 1$, $\Theta_{m,T} = 1$ and $R = 1 + r$. In (33), $f_{m,t+1}^*$ should be chosen such that it satisfies:

$$F_t = \frac{E_t \left[\exp \left\{ a_m R^{T-(t+1)} (\alpha_{m,t+1}^\top W_{t+1,s} + f_{m,t+1}^* F_{t+1}) \Theta_{m,t} F_{t+1} \right\} \right]}{E_t \left[\exp \left\{ a_m R^{T-(t+1)} (\alpha_{m,t+1}^\top W_{t+1,s} + f_{m,t+1}^* F_{t+1}) \Theta_{m,t} \right\} \right] R}, \tag{34}$$

We obtained (34) by maximising $J_{m,t}(V_{m,t}, \alpha_{m,t+1:T}, f_{m,t+1:T}, \mathcal{W}_t)$ with respect to $f_{m,t+1:T}$ in each t as in Section 2.2.

Note, that now the reverse supply for the s th WD also depends on the position in the alternative financial investment. Thus, any changes in its price process will influence investor's supply for the s th WD.

3. Pricing Weather Derivatives Using Weather Data

In this section, we show an example on how to price rainfall derivatives using historical weather data from China and discuss other applications of the pricing approach presented.

3.1. Pricing Chinese Rain

Chinese farmers are exposed to pronounced weather risks as Turvey and Kong (2010). According to The World Bank (2007) the existing agricultural insurance schemes are too expensive for Chinese agricultural producers. One of the causes might be the fact that crop insurer are exposed to variability in precipitation due to the impact of the later on farmers' crop production. Trading WDs can play an important role in transferring part of the weather exposure to financial markets and so

make crop insurance affordable for farmers. Since the weather exposed income of such a crop insurer is generally dependent on the rainfall in a number of geographical sites where the insured farmers are located, a basket of rainfall derivatives on the relevant sites should be used for hedging the risks. Motivated by this example, we illustrate pricing a basket of rainfall options on two weather stations Changde and Enshi¹ located in an agricultural area of China.

3.1.1. Setup

Our representative buyer is an insurer who offers crop insurance to farmers in the agricultural area of Changde and Enshi is, therefore, exposed to fluctuations in the local rainfall amount. The company discovers that its income is highly dependent on the cumulative precipitation in Changde and Enshi during May, and wants to hedge its rainfall exposure by holding a portfolio of put options on cumulative rainfall in these sites. Suppose, there is an investor who wants to issue such rainfall derivatives on the two of the relevant sites: Changde and Enshi. Then, the parties consider pricing a basket of derivatives containing two put options on cumulative precipitation over May in Changde and Enshi. They construct the put options as plain vanilla options on the underlying rainfall indices computed as the sums of monthly rainfall over May measured in each of the sites. Both options mature at the end of the measurement period, that is, on the 31st of May. We also suppose that the parties agree to renegotiate the put prices in the middle of the measurement period: on the 15th of May.

Let $\mathcal{R}_{z,i}$ denote the rainfall measured in the station i in day z . The cumulative precipitation over a period Z is then $\sum_{z \in Z} \mathcal{R}_{z,i}$. Thus, a put option on the cumulative precipitation in Z in station i with strike K_i (in precipitation units) has the payoff:

$$k \max\{K_i - \sum_{z \in Z} \mathcal{R}_{z,i}, 0\} \tag{35}$$

where k specifies the tick value of the option in monetary units.

Suppose, the insurer's income I_1 exhibits a non-linear dependence to the rainfall index in May of the form $I_1 \sim \log \mathcal{N}\{\mu_I(R_{z,i}), \sigma_I(R_{z,i})\}$ with $\mu_I(R_{z,i}) = 10 + \prod_{i \in 1,2} \sum_{z \in \text{May}} \mathcal{R}_{z,i} / K_i$ and $\sigma_I(R_{z,i}) = 0.1\{1 + \sum_{i \in 1,2} (1 - \mathcal{R}_{z,i} / K_i)^2\}$ parametrize the dependence of I_1 to the rainfall of Changde (indexed by 1) and Enshi (indexed by 2) respectively. Further parameters to specify are: the risk aversion $a_j = a_m = 10^{-4}$, the strikes $K_1 = K_2 = 1000$ and the tick value $k = 1$ monetary units per mm precipitation measured.

The log-return on the alternative investment in the investor's portfolio is assumed to follow a normal distribution with zero mean and annual volatility $\sigma_F = 0.1$ in the low volatility scenario and $\sigma_F = 0.25$ in the high volatility scenario. The value of the alternative investment in $t = 0$ is normalized to a hundred monetary units, that is $F_0 = 100$.

We compute the prices for the put options assuming different investment horizons T . If $T = 1$, in $t = 0$ (prior to the 1st of May) the agents negotiate the prices for the specified options under their current states of the world $\mathcal{W}_{1,0}$ and $\mathcal{W}_{2,0}$ respectively, and at $t = T$ (here, on 31 May) the settlement occurs according to the terminal time with realized payoffs W_T . $T = 2$ means that the agents additionally renegotiate the prices at $t = 1$ (here, on 15 May). The renegotiation at $t = 1$ takes place under the new circumstances arising from their new states $\mathcal{W}_{1,1}$ and $\mathcal{W}_{2,1}$ respectively. The final payments are then settled according to the realizations of W_2 . In our example, the state $\mathcal{W}_{i,t}$, $i = 1, 2$ includes the evolution of the rainfall, the observed price process $W_{n,n \leq t}$, F_t , λ_t , p_t , and r . The evolution to the next state $\mathcal{W}_{i,t+1}$ conditional on the realization of $\mathcal{W}_{i,t}$ is described by continuous function $g_{i,t}$. The terminal state of the insurer $\mathcal{W}_{1,T}$ contains also her realised income I_1 .

At $t \leq T$ denote the index value underlying the payoff in (35) in station i as $RX_{t,i}$. Then, $RX_{0,i} = 0$, $RX_{T,i} = \sum_{z \in 1.-31.May} \mathcal{R}_{z,i}$, and by $T = 2$ $RX_{1,i} = \sum_{z \in 1.-15.May} \mathcal{R}_{z,i}$. The corresponding put option

¹ Station numbers given by the World Meteorological Organisation are 57662 for Changde and 57447 for Enshi.

price, denoted as $W_{t,i}$, corresponds in $t = T$ to the put payoff, $W_{T,i} = \max(K_i - RX_{T,i}, 0)$, and for $t < T$ $W_{t,i}$ it is defined by the intersection of buyer's demand (17) and investor's supply (14).

3.1.2. Generation of Dependent Rainfall Paths on a Daily Basis

We use a model for multi-site rainfall generation proposed by Wilks (1998). This model accounts for spatial correlation of rainfall occurrences and rainfall amounts, and is flexible enough to describe different rainfall-based indexes. Based on Wilks (1998) we give the basic ideas of the model below:

Precipitation at time t in station s' $\mathcal{R}_{s',t}$ is modelled as:

$$\mathcal{R}_{s',t} = r_{s',t}X_{s',t} \tag{36}$$

where $X_{s',t}$ takes values $\{0,1\}$ and represents a rainfall occurrence at time t in station s' and $r_{s',t}$ is a positive random variable for the rainfall amount. This structure ensures that $\mathcal{R}_{s',t}$ equals zero whenever $X_{s',t}$ is zero and there is no rain and $\mathcal{R}_{s',t}$ equals $r_{s',t}$ whenever rainfall occurs. $X_{s',t}$ are assumed to form a Markov chain with two states, wet and dry:

$$X_{s',t} = \begin{cases} 1 \text{ (wet, } \geq X_{min}), \\ 0 \text{ (dry, } < X_{min}). \end{cases} \tag{37}$$

We take a first order Markov model where the probability of a wet day depends only on the state in the previous day. Our justification of this choice is based on the Bayesian information criteria (BIC), see Katz (1981) where the lowest value of the test statistics indicates the appropriate order of the Markov chain. The transition probabilities to the wet state are:

$$P(X_{s',t} = 1 | X_{s',t-1} = 0) = p_{01,s',t} \tag{38}$$

$$P(X_{s',t} = 1 | X_{s',t-1} = 1) = p_{11,s',t}. \tag{39}$$

The multi-site feature is added to the model through the generation of correlated occurrences in neighbour locations. We define the threshold probability $p_{s',crit}$:

$$p_{crit,s',t} = \begin{cases} p_{01,s',t} \text{ if } X_{s',t-1} = 0, \\ p_{11,s',t} \text{ if } X_{s',t-1} = 1, \end{cases} \tag{40}$$

and $X_{s',t}$ can be generated using

$$X_{s',t} = \begin{cases} 1 \text{ if } w_{s',t} \leq \Phi^{-1}(p_{crit,s',t}), \\ 0 \text{ if } w_{s',t} > \Phi^{-1}(p_{crit,s',t}). \end{cases} \tag{41}$$

Here $\Phi(\cdot)$ is the cumulative distribution function of the standard normal distribution, $\{w_{s,t}\}_{s \in \mathcal{S}} \sim N(0_{|\mathcal{S}|}, \Sigma)$, with $\Sigma_{s,s'} = \text{Corr}(w_{s,t}, w_{s',t})$ such that the empirical correlations $\text{Corr}(X_{s,t}, X_{s',t})$ of the rainfall occurrences are mimicked in the generated rainfall occurrence series, see Wilks (1998) for further details.

The multi-site rainfall amount, conditioned on a rainy day, $r_{s,t} | X_{s,t} = 1$ follows a mixture of two exponential distributions with a time dependent mixing parameter $\gamma_{s,t}$ and time changing means $\beta_{1,s,t}, \beta_{2,s,t}$. Following Wilks (1998) we can generate rainfall amounts at time t in site s using

$$r_{s,t} = r_{min} - \beta_{s,t} \log \Phi(v_{s,t}) \tag{42}$$

where

$$\beta_{s,t} = \begin{cases} \beta_{1,s,t} \text{ if } \Phi(w_{s,t}) / p_{s,crit} \leq \gamma_{s,t}, \\ \beta_{2,s,t} \text{ if } \Phi(w_{s,t}) / p_{s,crit} > \gamma_{s,t}, \end{cases} \tag{43}$$

and $v_{s,t}$ are normal covariates correlated such that the generated rainfall time series mimic the empirical correlation in the rainfall data.

Now we estimate the parameters for the multi-site rainfall generation in Changde and Enshi based on the rainfall data. Table 1 summarises the daily rainfall data for Changde and Enshi acquired via Research Data Centre of CRC 649 (Collaborative Research Centre 649: Economic risk).

Table 1. Description of the rainfall data.

Station	Number	Latitude	Longitude	Start Date	End Date
Changde	57,662	29.05	111.68	1951/01/01	2009/11/30
Enshi	57,447	30.28	109.47	1951/08/01	2009/11/30

BIC in Table 2 indicates that the first order Markov chain is appropriate to model the rainfall occurrences in the data.

The parameters for the rainfall amounts were estimated using maximum likelihood. For the analysed rainfall data $p_{01,1,t \in May} = 0.39$ and $p_{11,1,t \in May} = 0.59$ whereas $p_{01,2,t \in May} = 0.43$ and $p_{11,2,t \in May} = 0.64$ where the indices 1 and 2 refer to Changde and Enshi respectively. The empirical counterpart of $\text{Corr}(X_{1,t \in May}, X_{2,t \in May})$ appears to be 0.53, in order to mimic this correlation $\text{Corr}(w_{1,t \in May}, w_{2,t \in May})$ must be set to 0.76. The empirical counterpart of $\text{Corr}(r_{1,t \in May}, r_{2,t \in May} | X_{1,t \in May} = 1, X_{2,t \in May} = 1)$ is 0.16 for the considered data, and in order to obtain this correlation in the generated rainfall series $\text{Corr}(v_{1,t \in May}, v_{2,t \in May})$ must be 0.25.

Table 2. BIC criterion for different orders of Markov model for the rainfall occurrences in May

Order/BIC	Changde	Enshi
0	70.83	60.02
1	53.21	43.21
2	53.47	44.69
3	65.64	59.72

The results on fitting the mixture of two exponential distributions to $r_{i,t \in May} | X_{i,t \in May} = 1, i = \{1, 2\}$ are presented in Table 3.

Table 3. Maximum Likelihood estimator for the mixture of two exponential distributions for the rainfall amounts.

Parameter	Changde	Enshi
$\gamma_{\cdot, t \in [\tau_1, \tau_2]}$	0.78	0.58
$\beta_{1, \cdot, t \in [\tau_1, \tau_2]}$	15.90	23.14
$\beta_{2, \cdot, t \in [\tau_1, \tau_2]}$	0.62	1.86

In our Monte Carlo example, we generate the correlated rainfall time series in Changde and Enshi using the estimated parameters and (36) 10^4 times.

3.1.3. Results

By altering T, p, r and σ_F we obtain the put prices on the rainfall in Changde and Enshi under different market scenarios, see Table 4. All prices are normalized to the price obtained for $T = 1, p = 0, r = 0.01$ and $\sigma_F = 0.1$ to enable the comparison.

The shift of the demand curve by changing probability to default and capital costs is shown in Figure 1.

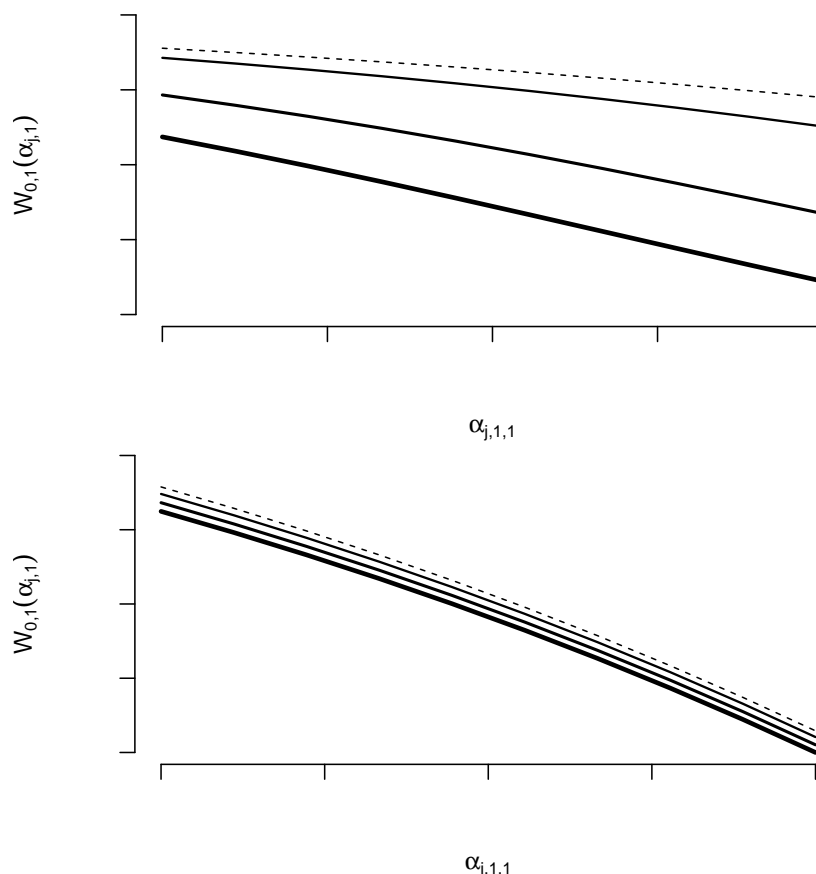


Figure 1. Shifts of the demand curves by increasing default risk (**top**) and capital costs (**bottom**). **Top:** the default risk probability $p_t = p$ changes from 0 (dashed) to 0.01, 0.05, 0.10 (solid lines, thicker with increasing p). **Bottom:** capital cost p.a. r changes from 1% (dashed) to 5%, 10%, 15% (solid lines, thicker with increasing r) for $\sigma_F = 0.1$. Both plots: the x -axis shows $\alpha_{j,1,1}$, the buyer’s position in the put option on the rainfall in Change during May ($CR_{May,1}$) with strike $K = 100$ and investment horizon $T = 1$; the y -axis shows $W_{0,1}(\alpha_{j,1})$, the price our representative buyer is willing to pay for such an option, where $\alpha_{j,1} = (\alpha_{j,1,1}, \alpha_{j,1,2})^\top$ and while $\alpha_{j,1,2}$, the position in the other put option on the rainfall in Enshi during May ($CR_{May,2}$) is kept constant.

The prices are significantly lower under a non-zero default probability and increasing capital costs. Whereas non-zero default probability has no effect on the supply of the WDs, the demand curve shifts downwards which results in lower equilibrium prices for the rainfall options. A higher risk-free rate r also results in lower WD prices to compensate for higher capital costs.

Figure 2 shows buyer’s demand in $t = 0$ for different time horizons T . In the “flexible” case $T = 2$ the WD prices are renegotiated at $t = 1$, and in the case $T = 1$ no rebalancing takes place. From Figure 2 we observe that buyer’s demand price elasticity at $t = 0$ is lower in the “flexible” case. Consistent with the classical result of Allen and Postlewaite (1984), already at $t = 0$ individual demand reflects all agents’ expectations, including the expectations about future trading behaviour.

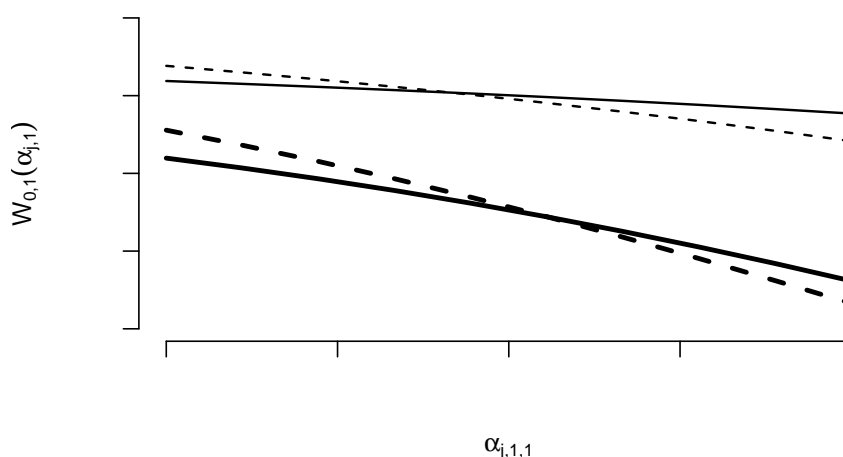


Figure 2. Shifts of the demand curves by increasing investment horizon T . Buyer’s demand with investment horizon $T = 1$ (dashed) and $T = 2$ (solid). Thinner lines correspond to the case capital costs $r = 5\%$ and probability of issuer’s default $p_t = p = 0$; thicker lines correspond to capital costs $r = 5\%$ and probability of issuer’s default $p_t = p = 0.05$ for $\sigma_F = 0.1$. The x -axis shows $\alpha_{j,1,1}$, the buyer’s position in the put option on the rainfall in Changde during May ($CR_{May,1}$) with strike $K = 100$; the y -axis shows $W_{0,1}(\alpha_{j,1})$, the price our representative buyer is willing to pay for such an option, where $\alpha_{j,1} = (\alpha_{j,1,1}, \alpha_{j,1,2})^\top$ and while $\alpha_{j,1,2}$, the position in the other put option on the rainfall in Enshi during May ($CR_{May,2}$) is kept constant.

Table 4. Put option prices on cumulative rainfall in Changde and Enshi under different market scenarios.

Scenarios	Put on $CR_{May,1}$		Put on $CR_{May,2}$		
	$T = 1$	$T = 2$	$T = 1$	$T = 2$	
$\sigma_F = 0.1$					
$r = 1\%$ $p = 0$	100.00	100.30	100.00	96.68	
$r = 5\%$ $p = 0$	99.67	99.95	98.81	96.35	
$r = 1\%$ $p = 0.05$	91.22	93.95	86.74	87.45	
$r = 5\%$ $p = 0.05$	90.87	93.62	85.95	87.12	
$\sigma_F = 0.25$					
$r = 1\%$ $p = 0$	100.00	100.31	100.23	96.68	
$r = 5\%$ $p = 0$	99.67	99.97	99.71	96.36	
$r = 1\%$ $p = 0.05$	91.23	94.23	86.88	87.73	
$r = 5\%$ $p = 0.05$	90.92	93.99	86.47	87.51	

Figure 3 shows investor’s supply in $t = 0$ for different time horizons T and for different levels in volatility of the alternative investment in the investor’s portfolio. In the “flexible” case $T = 2$ the WD prices are renegotiated at $t = 1$, and in the case $T = 1$ no rebalancing takes place. From Figure 3 we observe that the price elasticity of the investor’s supply at $t = 0$ is lower in the “flexible” case. Moreover, the reaction of the supply to the changes in market volatility is very subtle in the “flexible” case (under 1% and is not visible in the graph), as the downward shift of the supply curve is quite pronounced for $T = 1$. In this case, increase in market volatility (or volatility of the alternative investment) significantly stimulates investor’s supply for WDs.

Our example shows pricing rainfall options. The presented pricing model, however, is not limited to any particular kind of weather derivatives. In fact, it is possible to price various weather derivatives with such underlying indices as snowfall, sunshine hours, number of sunny days, number of rainy days, wind speed, and other weather indices of a practical relevance for retail, tourism and renewable energy operators.

Along with the other assumptions, the premise is the existence of a probabilistic model that precisely enough describes the distribution and the time evolution of the underlying weather index.

Then, the conditional expectations of agents' utilities, which determine their supply and demand, can be approximated using Monte Carlo samples from the fitted probabilistic model. Applying partial market clearing condition, one obtains the equilibrium prices for weather derivatives.

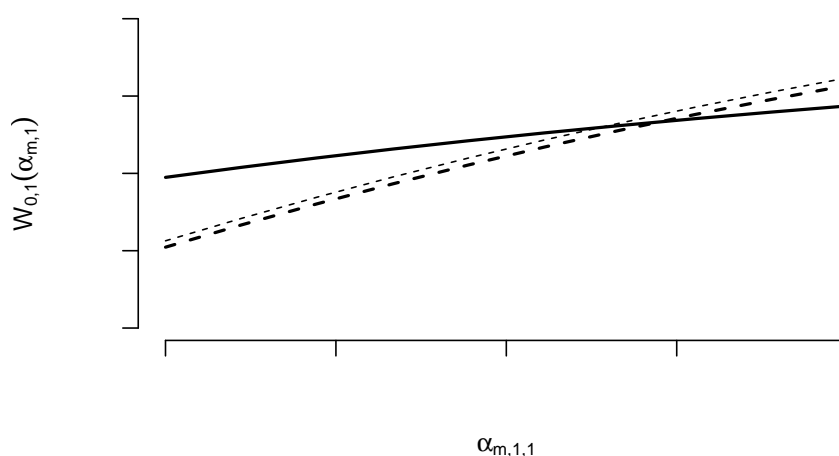


Figure 3. Shifts of the supply curves by increasing volatility of the alternative investment. Investor's supply with investment horizon $T = 1$ (dashed) and $T = 2$ (solid). Thinner lines correspond to the case low volatility of the alternative investment; thicker lines correspond to high volatility. The x -axis shows $\alpha_{m,1,1}$, the investor's position in the put option on the rainfall in Changde during May ($CR_{May,1}$) with strike $K = 100$; the y -axis shows $W_{0,1}(\alpha_{m,1})$, the price for which the investor is willing to sell such an option, where $\alpha_{m,1} = (\alpha_{m,1,1}, \alpha_{m,1,2})^\top$ and while $\alpha_{m,1,2}$, the position in the other put option on the rainfall in Enshi during April and May ($CR_{May,2}$) is kept constant.

4. Summary

We derive a dynamic utility-based model for pricing baskets of weather derivatives on multiple dependent underlying indices. Using dynamic programming approach to portfolio optimisation over a finite investment horizon and partial market clearing, we obtain semi-closed forms for the equilibrium prices of weather derivatives and for the optimal trading strategies. We provide extensions of the model to account for counterparty default risk and a possibility of an alternative financial investment.

As expected, there is an adverse effect of increasing default risk and capital costs on the demand for weather derivatives and on their prices. We find, however, a stimulating effect of increasing market volatility on the supply for weather derivatives.

We apply the proposed model to price rainfall options using historical rainfall data of agricultural provinces Changde and Enshi in China. The equilibrium put option prices on cumulative rainfall over May are compared for different market scenarios. The effects of increasing default risk, capital costs, investment horizon, and market volatility are assessed.

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Appendix A

Proof of Proposition 1. We prove the assertion by backward induction starting with $t = T - 1$. Utility maximisation problem of buyer j in $T - 1$:

$$\begin{aligned} \max_{\alpha_{j,T} \in \mathbb{R}^S} & J_{j,T-1}(V_{j,T-1}, \alpha_{j,T}, \mathcal{W}_{j,T-1}) \\ \text{s.t.} & \alpha_{j,T}^\top W_{T-1} + \beta_{j,T} B_{T-1} - V_{j,T-1} = 0. \end{aligned} \tag{A1}$$

The expected utility in $T - 1$ reads:

$$J_{j,T-1}(V_{j,T-1}, \alpha_{j,T}, \mathcal{W}_{j,T-1}) = E_{T-1} \{ -\exp(-a_j \Pi_{j,T}) \}. \tag{A2}$$

By plugging the self-financing constraint in (A2) we obtain

$$J_{j,T-1}(V_{j,T-1}, \alpha_{j,T}, \mathcal{W}_{j,T-1}) = E_{T-1} [-\exp\{-a_j(I_j + \alpha_{j,T}^\top W_T + R V_{j,T-1} - R \alpha_{j,T}^\top W_{T-1})\}]. \tag{A3}$$

Taking the first derivative of (A3) with respect to $\alpha_{j,T}$, we win the gradient vector with the s th entry equal to:

$$\begin{aligned} \frac{\partial J_{j,T-1}(V_{j,T-1}, \alpha_{j,T}, \mathcal{W}_{j,T-1})}{\partial \alpha_{j,T,s}} &= a_j \exp\{-a_j(V_{j,T-1} - \alpha_{j,T}^\top W_{T-1})R\} \\ &\cdot E_{T-1} [\exp\{-a_j(I_j + \alpha_{j,T}^\top W_T)\} (W_{T,s} - R W_{T-1,s})] \end{aligned} \tag{A4}$$

for $s = 1, \dots, S$. Taking the second derivative of (A3) with respect to $\alpha_{j,T}$ gives the Hessian matrix H with ss' th entry:

$$\begin{aligned} \frac{\partial^2 J_{j,T-1}(V_{j,T-1}, \alpha_{j,T}, \mathcal{W}_{j,T-1})}{\partial \alpha_{j,T,s} \partial \alpha_{j,T,s'}} &= -a_j^2 \exp\{-a_j(V_{j,T-1} - \alpha_{j,T}^\top W_{T-1})R\} \\ &\cdot E_{T-1} [\exp\{-a_j(I_j + \alpha_{j,T}^\top W_T)\} (W_{T,s} - R W_{T-1,s})(W_{T,s'} - R W_{T-1,s'})] \end{aligned} \tag{A5}$$

for $s, s' = 1, \dots, S$. The Hessian H defined entry-wise by (A5) is negative semi-definite, since for any nonzero $x \in \mathbb{R}^S$:

$$\begin{aligned} x^\top H x &= -a_j^2 \exp\{-a_j(V_{j,T-1} - \alpha_{j,T}^\top W_{T-1})R\} \\ &\cdot E_{T-1} [\exp\{-a_j(I_j + \alpha_{j,T}^\top W_T)\} \{ \sum_{s=1}^S x_s (W_{T,s} - R W_{T-1,s}) \}^2] \end{aligned} \tag{A6}$$

is always smaller or equal to zero. Hence, we obtain the maximiser of (A3) by setting the gradient (A4) coordinate-wise to zero, in particular:

$$E_{T-1} [\exp\{-a_j(I_j + \alpha_{j,T}^\top W_T)\} (W_{T,s} - R W_{T-1,s})] = 0. \tag{A7}$$

In $T - 1$ the current prices W_{T-1} are known and can be taken out of the expectation in (A7). Thus, we obtain the reverse demand of the j th buyer for each $s \in \mathcal{S}$ as:

$$W_{T-1,s}^j(\alpha_{j,T}) = \frac{E_{T-1} [\exp\{-a_j(I_j + \alpha_{j,T}^\top W_T)\} W_{T,s}]}{E_{T-1} [\exp\{-a_j(I_j + \alpha_{j,T}^\top W_T)\}] R} = \frac{E_{T-1} \{ \exp(-a_j \alpha_{j,T}^\top W_T) \Theta_{j,T} W_{T,s} \}}{E_{T-1} \{ \exp(-a_j \alpha_{j,T}^\top W_T) \Theta_{j,T} \} R} \tag{A8}$$

with $\Theta_{j,T} = \exp(-a_j I_j)$.

Applying partial market clearing or zero-net-supply condition to all WDs, we find $\alpha_{j,T}^*$ and $W_{T-1} = W_{T-1}^m(\alpha_{m,T}^*)$. Then, the maximised utility of buyer j at $T - 1$ is:

$$\begin{aligned} J_{j,T-1}^*(V_{j,T-1}, \mathcal{W}_{j,T-1}) &= J\{V_{j,T-1}, \alpha_{j,T}^*, W_{T-1}, \mathcal{W}_{j,T-1}\} \\ &= -\exp\{-a_j V_{j,T-1} R\} \Theta_{j,T-1} \end{aligned} \tag{A9}$$

with

$$\Theta_{j,T-1} = \exp\{a_j \alpha_{j,T}^{*\top} W_{T-1} R\} E_{T-1} [\exp\{-a_j (\alpha_{j,T}^{*\top} W_T)\} \Theta_{j,T}]. \tag{A10}$$

$J_{j,T-1}^*(V_{j,T-1}, \mathcal{W}_{j,T-1})$ is of the exponential form like the utility function itself, and the induction hypothesis holds for $t = T - 1$. Assume, that it holds for $t = T, T - 1, \dots, k$. Then, in $k - 1$:

$$\begin{aligned} \max_{\alpha_{j,k} \in \mathbb{R}^S} J_{j,k-1} \{V_{j,k-1}, \alpha_{j,k}, \mathcal{W}_{j,k-1}\} \\ \text{s.t. } \alpha_{j,k}^\top W_{k-1,s} + \beta_{j,k} B_{k-1} - V_{j,k} = 0. \end{aligned} \tag{A11}$$

is the constrained utility maximisation problem faced by buyer j . The expected utility in $k - 1$ is:

$$\begin{aligned} J_{j,k-1} \{V_{j,k-1}, \alpha_{j,k}, \mathcal{W}_{j,k-1}\} &= E_{k-1} \left\{ J_{j,k}^*(V_{j,k}, \mathcal{W}_{j,k}) \right\} \\ &= E_{k-1} \left\{ -\exp(-a_j V_{j,k} R^{T-k}) \Theta_{j,k} \right\}, \end{aligned} \tag{A12}$$

where

$$\Theta_{j,k} = \exp(a_j R^{T-k} \alpha_{j,k+1}^{*\top} W_k) E_k \{ \exp(-a_j R^{T-(k+1)} \alpha_{j,k+1}^{*\top} W_{k+1,s}) \Theta_{j,k+1} \}. \tag{A13}$$

Now we use the following identity:

$$V_{j,k} = R(V_{j,k-1} - \alpha_{j,k}^\top W_{k-1}) + \alpha_{j,k}^\top W_k \tag{A14}$$

to rewrite (A12) as a function of $\alpha_{j,k}$:

$$J_{j,k-1}(V_{j,k-1}, \alpha_{j,k}, \mathcal{W}_{j,k-1}) = E_{k-1} \left(-\exp[-a_j \{R(V_{j,k-1} - \alpha_{j,k}^\top W_{k-1}) + \alpha_{j,k}^\top W_k\} R^{T-k}] \Theta_{j,k} \right). \tag{A15}$$

By taking the derivative of (A15) with respect to $\alpha_{j,k}$ we find the gradient with sth entry:

$$\begin{aligned} \frac{\partial J_{j,k-1}(V_{j,k-1}, \alpha_{j,k}, \mathcal{W}_{j,k-1})}{\partial \alpha_{j,k,s}} &= a_j R^{T-k} \exp\{-a_j (V_{j,k-1} - \alpha_{j,k}^\top W_{k-1}) R^{T-k+1}\} \\ &\cdot E_{k-1} \left\{ -\exp(-a_j \alpha_{j,k}^\top W_k R^{T-k}) \Theta_{j,k}(W_{k,s} - R W_{k-1,s}) \right\} \end{aligned} \tag{A16}$$

As in (A6) the Hessian is also negative semi-definite. Thus, the maximiser of (A15) is found by setting the gradient (A16) to zero, that is:

$$E_{k-1} \left\{ -\exp(-a_j \alpha_{j,k}^\top W_k R^{T-k}) \Theta_{j,k}(W_{k,s} - R W_{k-1,s}) \right\} = 0. \tag{A17}$$

We obtain the reverse demand of the j th buyer for each $s \in \mathcal{S}$ as:

$$W_{k-1,s}^j(\alpha_{j,k}) = \frac{E_{k-1} \left\{ \exp(-a_j \alpha_{j,k}^\top W_k R^{T-k}) \Theta_{j,k} W_{k,s} \right\}}{E_{k-1} \left\{ \exp(-a_j \alpha_{j,k}^\top W_k R^{T-k}) \Theta_{j,k} \right\} R} \tag{A18}$$

Partial market clearing in $k - 1$ determines $\alpha_{j,k}^*, W_{k-1} = W_{k-1}^m(\alpha_{m,k}^*)$ and the maximised utility of buyer j in this period:

$$\begin{aligned} J_{j,k-1}^*(V_{j,k-1}, \mathcal{W}_{j,k-1}) &= J_{j,k-1} \{V_{j,k-1}, \alpha_{j,k}^*, \mathcal{W}_{j,k-1}\} \\ &= -\exp\{-a_j V_{j,k-1} R\} \Theta_{j,k-1} \end{aligned} \tag{A19}$$

with

$$\Theta_{j,k-1} = \exp\{a_j a_{j,k}^{*\top} W_{k-1} R\} E_{k-1} [\exp\{-a_j (a_{j,k}^{*\top} W_k)\} \Theta_{j,k}], \quad (\text{A20})$$

which completes the proof. \square

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