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ABSTRACT

Machine Learning Estimation of Heterogeneous Causal Effects: Empirical Monte Carlo Evidence*

We investigate the finite sample performance of causal machine learning estimators for heterogeneous causal effects at different aggregation levels. We employ an Empirical Monte Carlo Study that relies on arguably realistic data generation processes (DGPs) based on actual data. We consider 24 different DGPs, eleven different causal machine learning estimators, and three aggregation levels of the estimated effects. In the main DGPs, we allow for selection into treatment based on a rich set of observable covariates. We provide evidence that the estimators can be categorized into three groups. The first group performs consistently well across all DGPs and aggregation levels. These estimators have multiple steps to account for the selection into the treatment and the outcome process. The second group shows competitive performance only for particular DGPs. The third group is clearly outperformed by the other estimators.

JEL Classification: C21

Keywords: causal machine learning, conditional average treatment effects,

selection-on-observables, random forest, causal forest, lasso

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1 Introduction

Economists and many other professionals are interested in causal effects of policies or interventions. This has triggered substantial advances in microeconometrics and statistics in understanding the identification and estimation of different average causal effects in the recent decades (see, e.g., Imbens & Wooldridge, 2009; Athey & Imbens, 2017, and references therein). However, in most applications it is also interesting to look beyond the average effects in order to understand how the causal effects vary with observable characteristics. For example, finding those individuals who benefit most from active labor market policies, promotion campaigns or medical treatments is important for the efficient allocation of public and private resources.

In recent years, methods for the systematic estimation of heterogeneous causal effects have been developed in different research disciplines. Those methods adapt standard machine learning methods to flexibly estimate heterogeneity along a potentially large number of covariates. The suggested estimators use regression trees (Su, Tsai, Wang, Nickerson, & Li, 2009; Athey & Imbens, 2016), Random Forests (Wager & Athey, 2018; Athey, Tibshirani, & Wager, 2018), the least absolute shrinkage and selection operator (Lasso) (Qian & Murphy, 2011; Tian, Alizadeh, Gentles, & Tibshirani, 2014; Chen, Tian, Cai, & Yu, 2017), support vector machines (Imai & Ratkovic, 2013), boosting (Powers et al., 2018), neural nets (Johansson, Shalit, & Sontag, 2016; Shalit, Johansson, & Sontag, 2016; Schwab, Linhardt, & Karlen, 2018) or Bayesian machine learning (Hill, 2011; Taddy, Gardner, Chen, & Draper, 2016). Recently, the first applied studies using these methods appeared in economics (e.g., Bertrand, Crépon, Marguerie, & Premand, 2017; Davis & Heller, 2017; Knaus, Lechner, & Strittmatter, 2017; Andini, Ciani, de Blasio, D'Ignazio, & Salvestrini, 2018; Ascarza, 2018; Strittmatter, 2018).

In contrast to the rather mature literature about the estimation of average causal effects, the literature on the estimation of effect heterogeneity is still lacking guidance for practitioners about which methods are well suited for their intended applications. Theoretical asymptotic approximations are currently either not available, incomplete, or

¹Hastie, Tibshirani, and Friedman (2009) introduce the underlying machine learning algorithms. Athey (2018) and Belloni, Chernozhukov, and Hansen (2014a) provide an overview how those methods might be used in the estimation of average causal effects and other parameters of interest.

they are based on non-overlapping assumptions preventing comparisons of estimators. Furthermore, the available information about the finite sample performance is of limited use to practitioners. Most comparisons are based on data generating processes (DGPs) that are very unrealistic for real applications. One exception is Wendling et al. (2018) who base their simulation study on data from medical records. However, they focus in their study on the special case of binary outcomes and data structures that are unusual in economics.

In this study, we categorize major approaches from different fields. We distinguish between generic approaches that can be combined with a variety of different off-the-shelf machine learning estimators and estimator specific approaches that modify an existing method. The generic approaches are combined with the machine learning estimators Random Forest and Lasso. This leads to eleven different causal machine learning estimators under investigation. As opposed to standard simulation methods that rely on a synthetic DGP, we investigate the finite sample performance of these estimators in an Empirical Monte Carlo (EMCS) approach (e.g., Huber, Lechner, & Wunsch, 2013; Lechner & Wunsch, 2013). An EMCS informs the DGPs as much as possible by real data and reduces synthetic components in the DGP to a minimum. We consider six different specifications of the heterogeneous causal effects, two different sample sizes, and DGPs with and without selection into treatment.²

Our contribution to the aforementioned literature is three-fold: First, we provide a comprehensive comparison of different estimators and DGPs. Second, we consider the finite sample properties of causal machine learning estimators for effect heterogeneity under DGPs that are arguably realistic at least in some fields of economics. Third, this is the first simulation study that considers also different aggregation levels of the heterogeneous effects. In particular, we consider an intermediate aggregation level between the most individualized causal effects and the average population effect. Such intermediate aggregation levels are important as feasible action rules for practitioners.

Our findings suggest that no causal machine learning estimator is superior for all DGPs and aggregation levels. However, four estimators show a relatively good performance in all

 $[\]overline{^{2}\text{We focus}}$ on point estimation and leave the investigation of inference methods for further research.

settings: Random Forests combined with a doubly robust outcome modification (based on Chernozhukov et al., 2018), Causal Forest with local centering (Athey et al., 2018), Lasso combined with a covariate modification and efficiency augmentation (Tian et al., 2014), and Lasso with R-learning (Nie & Wager, 2017). All those methods use multiple estimation steps to account for the selection into treatment and the outcome process. Several other estimators may be suitable in specific empirical settings but their performance is unstable across different DGPs. Lasso estimators tend to be more unstable than Random Forests, which frequently prevents them from achieving a normal distribution.

In the next section, we introduce the notation and the estimation targets. In Section 3, we describe and categorize causal machine learning approaches to estimate heterogeneous causal effects. In Section 4, we explain the implementation of the estimators. In Section 5, we discuss the EMCS approach. In Section 6, we provide our simulation results. The final Section concludes and hints at some topics for future research. Appendices A-D contain supplementary statistics and results. Appendix A describes the data that are used for the EMCS. Appendices B and C provide details about the DGPs and the implementation of estimators, respectively. Finally, Appendix D shows and discusses the full simulation results of all DGPs. We provide code that implements the estimators under investigation in the R package CATEs on GitHub.

2 Notation and estimation targets

We describe the parameters of interest using Rubin's (1974) potential outcome framework. The dummy variable D_i indicates a binary treatment, e.g. participation in a training program. Let Y_i^1 denote the outcome (e.g., employment) if individual i (i = 1, ..., N) receives the treatment ($D_i = 1$). Correspondingly, Y_i^0 denotes the outcome if individual i does not receive the treatment ($D_i = 0$). Each individual can either receive the treatment or not. This means that only one of the two potential outcomes (Y_i^d) is observable:

$$Y_i = D_i Y_i^1 + (1 - D_i) Y_i^0. (1)$$

Thus, the individual treatment effect (ITE) $\xi_i = Y_i^1 - Y_i^0$ of D_i on Y_i is never observed.

However, the identification of expectations of ξ_i may be possible under plausible assumptions. For example, the identification of the average treatment effect (ATE), $\tau = E[\xi_i]$, or the average treatment effect on the treated (ATET), $\theta = E[\xi_i \mid D_i = 1]$, are standard in microeconometrics (see, e.g., Imbens & Wooldridge, 2009).

The focus of this study is on conditional average treatment effects (CATEs). CATEs take the expectations of ξ_i conditional on exogenous pre-treatment covariates.³ We call the finest conditioning level that uses all available covariates X_i individualized average treatment effect (IATE),

$$\tau(x) = E[\xi_i \mid X_i = x] = \mu^1(x) - \mu^0(x), \tag{2}$$

where $\mu^d(x) = E[Y_i^d \mid X_i = x]$ denotes the conditional expectation of the unobserved potential outcomes. IATEs provide an approximation of ITEs for the set of covariates that are at the disposal of the researcher in a specific application. However, researchers may additionally be interested in intermediate aggregation levels that are coarser than IATEs but finer than ATEs. Especially groups based on a smaller set of pre-defined characteristics, G_i , may be of interest if the estimated IATEs need to be summarized for the research community, communicated to practitioners, or acted upon.⁴ We call the effects defined on this aggregation level group average treatment effects (GATE),

$$\tau(g) = E[\xi_i \mid G_i = g] = \int \tau(x) f_{X_i \mid G_i = g}(x) dx.$$
 (3)

The identification of any aggregation level of individual treatment effects in observational studies is complicated by non-random treatment assignment. However, identification of the IATE and any coarser aggregation level is still possible if the observable covariates X_i contain all confounders.⁵ These are covariates that jointly affect the treatment probability and the potential outcomes. Although there are alternative ways to identify the various effects, here we focus on the case where all the confounders are contained in the data available to the researcher. This means that we operate throughout the paper under the

³Covariates are also called features or predictors in parts of the machine learning literature.

⁴For example, if interest is in gender differences, $G_i \in \{female, male\}$.

 $^{{}^{5}}X_{i}$ represents the union of confounders and heterogeneity variables for notational convenience. In principle, they may be completely, partly or non-overlapping (see, e.g., Knaus et al., 2017).

following assumptions.

Assumption 1: (Conditional independence): $Y_i^1, Y_i^0 \perp \!\!\! \perp D_i \mid X_i = x$, for all x in the support of X_i .

Assumption 2: (Common support): $0 < P[D_i = 1 \mid X_i = x] = p(x) < 1$, for all x in the support of X_i .

Assumption 3: (Exogeneity of covariates): $X_i^1 = X_i^0.6$

Assumption 4: (Stable Unit Treatment Value Assumption, SUTVA): $Y_i = Y_i^1 D_i + Y_i^0 (1 - D_i)$.

Assumption 1 states that the potential outcomes are independent of the treatment conditional on the confounding covariates. According to Assumption 2, the conditional treatment probability (often called propensity score) is bounded away from zero and one. Assumption 3 requires that the covariates are not affected by the treatment. Assumption 4 excludes spillover effects between treated and non-treated. Under Assumptions 1-4,

$$E[Y_i^d \mid X_i = x, D_i = 1 - d] = E[Y_i \mid X_i = x, D_i = d] = \mu(d, x)$$
(4)

$$\Rightarrow \tau(x) = \mu(1, x) - \mu(0, x), \tag{5}$$

and thus IATEs, GATEs and ATE are identified from observable data. We denote the conditional expectations of the outcomes in one treatment arm by $\mu(d, x) = E[Y_i \mid X_i = x, D_i = d]$, the conditional expectation of the outcome as $\mu(x) = E[Y_i \mid X_i = x]$, and the conditional treatment probability by $p(x) = P[D_i = 1 \mid X_i = x]$.

3 Causal machine learning of effect heterogeneity

Equation 5 shows that the fundamental task is to estimate the difference of two conditional expectations. However, we never observe the differences at the individual level and have to estimate them in two different subpopulations. Thus, the estimation of IATEs is a non-standard machine learning problem. In this section, we present different approaches to target the estimation of IATEs. We distinguish between generic approaches and one estimator specific approach. Generic approaches split the causal estimation problem

 $[\]overline{^6{\rm The~potential~confounders~}X_i^d{\rm~are~defined}}$ equivalently to potential outcomes.

into several standard prediction problems and may be combined with a large variety of supervised machine learning estimators. On the other hand, Causal Forest Athey et al. (2018) is a modification of a specific machine learning estimator to move the target from the estimation of outcomes to the estimation of IATEs.

3.1 Generic approaches

A straightforward generic approach follows directly from Equation 5. Conditional mean regressions takes the difference of conditional expectations that are estimated in the two samples of treated and non-treated separately using off-the-shelf machine learning methods to estimate the conditional outcome means $\hat{\mu}(d, x)$:⁷

$$\hat{\tau}_{CMR}(x) = \hat{\mu}(1, x) - \hat{\mu}(0, x). \tag{6}$$

Conditional mean regressions are straightforward to implement. Any supervised machine learning methods for conditional mean estimation may be used. However, their usual target is to minimize the mean squared error (MSE) in two separate prediction problems and they are not tailored to estimate IATEs.⁸ This suggests that they may be outperformed by more specialized methods for this causal problem. Three generic multi-step approaches that target IATE estimation are presented in the following and a framework to summarize them is provided.

3.1.1 Modified outcome methods

Abadie (2005) introduces the idea of modifying the outcome to estimate conditional average treatment effects on the treated in studies based on difference-in differences. For IATEs in the experimental and observational setting, the idea is formulated by Signorovitch (2007) and Zhang, Tsiatis, Laber, and Davidian (2012), respectively. The latter discuss two modifications of the outcome, which we summarize as modified outcome methods (MOM).

⁷This approach is also referred to as T-Learner (Künzel, Sekhon, Bickel, & Yu, 2017; Nie & Wager, 2017) or Q-Learning (Qian & Murphy, 2011).

 $^{^8}$ For an intuition why this is not optimal: Biases that for the same value of x go in the same direction are less harmful than if they go in opposite directions. However, this cannot be accounted for by separate MSEs that are not directly linked (for a new Causal Forest estimator that takes up this theme directly, see Lechner (2018)).

The first is based on inverse probability weighting (IPW) (e.g., Horvitz & Thompson, 1952; Hirano, Imbens, & Ridder, 2003) where the modified outcome is

$$Y_{i,IPW}^* = Y_i \frac{D_i - p(X_i)}{p(X_i)(1 - p(X_i))}. (7)$$

The second is based on the doubly robust (DR) estimator of Robins and Rotnitzky (1995),

$$Y_{i,DR}^* = \mu(1, X_i) - \mu(0, X_i) + \frac{D_i(Y_i - \mu(1, X_i))}{p(X_i)} - \frac{(1 - D_i)(Y_i - \mu(0, X_i))}{(1 - p(X_i))}.$$
 (8)

The crucial insight here is that $\tau(x) = E[Y_{i,IPW}^* \mid X_i = x] = E[Y_{i,DR}^* \mid X_i = x]$. This means that a regression with one of these modified outcomes and covariates X_i can be used to obtain estimates of IATEs, $\hat{\tau}_{IPW}(x)$ or $\hat{\tau}_{DR}(x)$. In practice, the researcher has no access to the true parameters p(x) and $\mu(d,x)$, the so-called nuisance parameters. The conditional expectations need to be approximated in a first step and plugged into Equations 7 and 8. Any suitable prediction method can be used to estimate the nuisance parameters as well as the IATEs.

The asymptotic properties of $E[Y_{i,DR}^*]$ as estimator for ATEs are well understood (Belloni, Chernozhukov, & Hansen, 2014b; Farrell, 2015; Belloni, Chernozhukov, Fernández-Val, & Hansen, 2017; Chernozhukov et al., 2017, 2018). Furthermore, Abrevaya, Hsu, and Lieli (2015) and Lee, Okui, and Whang (2017) analyze the properties of estimating $\tau(z) = E[Y_{i,DR}^* \mid Z_i = z]$ for a low-dimensional subset of covariates (Z_i) , respectively. However, both do not consider machine learning to estimate the nuisance parameters. We are currently not aware of theoretical results for the case where nuisance parameters and IATEs are estimated with machine learning.

Simulation evidence of Powers et al. (2018) suggests that estimators based on $Y_{i,IPW}^*$ may exhibit high variance due to potentially extreme values of the propensity score in the denominator. Estimators based on $Y_{i,DR}^*$ might be more stable, because of the double-robustness property, but this is unexplored until now.

3.1.2 Modified Covariate Method

Tian et al. (2014) introduce the modified covariate method (MCM) for experiments and Chen et al. (2017) extend it to observational studies. They show that we can estimate IATEs by solving the objective function

$$\min_{\tau} \left[\frac{1}{N} \sum_{i=1}^{N} T_i \frac{D_i - p(X_i)}{p(X_i)(1 - p(X_i))} \left(Y_i - \frac{T_i}{2} \tau(X_i) \right)^2 \right], \tag{9}$$

where $T_i = 2D_i - 1 \in \{-1, 1\}$. The name MCM results from the practice to replace the non-parametric function of the IATE with a linear working model, $\tau(x) = x\beta$. This enables us to rewrite the minimization problem 9 as

$$\hat{\beta}_{MCM} = \arg\min_{\beta} \left[\frac{1}{N} \sum_{i=1}^{N} T_i \frac{D_i - p(X_i)}{p(X_i)(1 - p(X_i))} \left(Y_i - X_i^{MCM} \beta \right)^2 \right], \tag{10}$$

where $X_i^{MCM} = T_i/2X_i$ are the modified covariates. The estimated IATEs are then obtained by $\hat{\tau}_{MCM}(x) = x\hat{\beta}_{MCM}$. The nuisance parameter p(x) needs to be estimated in a first step using any suitable method.

In principle, rewriting 9 as

$$\min_{\tau} \left[\frac{1}{N} \sum_{i=1}^{N} T_i \frac{D_i - p(X_i)}{4p(X_i)(1 - p(X_i))} \left(2T_i Y_i - \tau(X_i) \right)^2 \right], \tag{11}$$

allows to apply any machine learning estimator that is able to solve weighted minimization problems. However, we are not aware of any study that notices and pursues this possibility.

MCM does not require to specify any model of the so-called main effects $\mu(x)$ or $\mu(d, x)$. However, Tian et al. (2014) describe that an estimate of $\mu(x)$ might be useful to increase efficiency. The efficiency augmented version replaces the outcome Y_i in Equations 9 to 11 by the residuals $Y_i - \mu(X_i)$. Thus, MCM with efficiency augmentation (EA) requires additionally to estimate the nuisance parameter $\mu(x)$ in the first step. Tian et al. (2014) show that MCM with a linear working model provides the best linear predictor of the potentially non-linear $\tau(x)$. However, we are not aware of any further theoretical analyses of the statistical properties of this approach.

3.1.3 R-learning

Nie and Wager (2017) propose R-learning that is based on the partially linear model of Robinson (1988). It is equivalent to MCM with EA for 50:50 randomization but solves otherwise a different minimization problem to estimate IATEs:⁹

$$\min_{\tau} \left\{ \frac{1}{N} \sum_{i=1}^{N} \left[(Y_i - \mu(X_i)) - (D_i - p(X_i)) \tau(X_i) \right]^2 \right\}$$
 (12)

Like for MCM, current implementations consider a linear working model for the IATE (Nie & Wager, 2017; Zhao, Small, & Ertefaie, 2017) and solve

$$\hat{\beta}_{RL} = \underset{\beta}{\operatorname{arg\,min}} \left\{ \frac{1}{N} \sum_{i=1}^{N} \left[(Y_i - \mu(X_i)) - X_i^{RL} \beta \right]^2 \right\},\tag{13}$$

where $X_i^{RL} = (D_i - p(X_i))X_i$ can be considered as an alternative way to modify covariates. The estimated IATEs are then obtained by $\hat{\tau}_{RL}(x) = x\hat{\beta}_{RL}$. Similar to MCM, 12 can be rewritten as

$$\min_{\tau} \left\{ \frac{1}{N} \sum_{i=1}^{N} (D_i - p(X_i))^2 \left[\frac{Y_i - \mu(X_i)}{D_i - p(X_i)} - \tau(X_i) \right]^2 \right\}$$
 (14)

and solved with any suitable method after estimating the nuisance parameters in a first step (see also Schuler, Baiocchi, Tibshirani, & Shah, 2018).

Nie and Wager (2017) show that R-learning can perform as good as an oracle estimator that knows the real nuisance parameters in the special case of solving 12 with penalized kernel regression. This result requires that the estimators of the nuisance parameters need to be fourth root consistent in the semiparametric case.

3.1.4 Summary of generic approaches to estimate IATEs

One goal of this paper is to structure approaches that estimate IATEs coming from different literatures. One way to put the approaches above on more common ground is by noting that they can all be considered as solving a weighted minimization problem with modified

⁹Murphy (2003) and Robins (2004) develop a similar method for optimal dynamic treatment regimes called A-Learning. It is adapted for the binary treatment case by Chen et al. (2017). We focus on R-Learning because it explicitly applies machine learning methods in all estimation steps.

Table 1: Summary of generic approaches to estimate IATEs

Approach	w_i	Y_i^*
MOM IPW	1	$Y_{i,IPW}^*$
MOM DR	1	$Y_{i,DR}^*$
MCM	$T_{i} \frac{D_{i} - p(X_{i})}{4p(X_{i})(1 - p(X_{i}))}$ $T_{i} \frac{D_{i} - p(X_{i})}{4p(X_{i})(1 - p(X_{i}))}$	$2T_iY_i$
MCM with EA	$T_i \frac{D_i - p(X_i)}{4p(X_i)(1 - p(X_i))}$	$2T_i(Y_i - \mu(X_i))$
R-Learning	$(D_i - p(X_i))^2$	$\frac{Y_i - \mu(X_i)}{D_i - p(X_i)}$

outcomes:

$$\min_{\tau} \left\{ \frac{1}{N} \sum_{i=1}^{N} w_i \left[Y_i^* - \tau(X_i) \right]^2 \right\}. \tag{15}$$

Table 1 summarizes the weights, w_i , and outcome modifications, Y_i^* , underlying the different approaches. This common representation is helpful to see and understand the differences of the methods. The two MOM methods require no additional weighting because $\tau(x) = E[Y_{i,IPW}^* \mid X_i = x] = E[Y_{i,DR}^* \mid X_i = x]$. For MCM, $\tau(x) = E[2T_iY_i \mid X_i = x] = E[2T_i(Y_i - \mu(X_i)) \mid X_i = x]$ in the special case of 50:50 randomization of the treatment (p(x) = 0.5). Any deviating assignment mechanism requires reweighting with IPW weights to control for the deviation from 50:50 randomization. The modified outcome of R-learning is equivalent to MCM with efficiency augmentation if p(x) = 0.5. However, while the MCM modified outcome does not change with other propensity scores and preserves the interpretation as mean comparison under 50:50 randomization, $E[(Y_i - \mu(X_i))/(D_i - p(X_i)) \mid X_i = x]$ is lacking such an intuitive interpretation.

3.2 Causal Forest

Another strand of literature modifies machine learning algorithms based on regression trees (Breiman, Friedman, Stone, & Olshen, 1984) to estimate IATEs. We focus on Causal Forest that is a special case of the Generalized Random Forest of Athey et al. (2018) as

the most recent estimator in this line of research.¹⁰ Causal Forests build on the idea of Random Forests (Breiman, 2001) and boil down to taking the difference of two weighted means in our case of binary treatments,

$$\hat{\tau}_{CF}(x) = \sum_{i=1}^{N} D_i w_i^1(x) Y_i - \sum_{i=1}^{N} (1 - D_i) w_i^0(x) Y_i, \tag{16}$$

where the weights $w_i^d(x)$ define an adaptive local neighbourhood around the covariate value of interest, x. These weights are obtained from the tailored splitting procedure that we describe in Section 4.1. Athey et al. (2018) show that this estimator can be consistent and asymptotically normal for a fixed covariate space.

4 Implementation of estimators

Section 3.1 describes generic approaches to split the estimation of IATEs into several prediction problems. The machine learning literature offers a large variety of potential methods such that the investigation of every possible combination of approaches and machine learning methods is not feasible given our restrictions on computational costs. Thus, this study is restricted to two machine learning methods for the implementation of the prediction problems. They are chosen to be representative for more general approaches. First, we consider Random Forests (Breiman, 2001) that serve as a popular representative for methods that attempt *local* approximations of conditional mean functions. Second, we consider Lasso (Tibshirani, 1996) as a method that attempts *global* approximation of conditional mean functions.¹¹ Both methods are increasingly popular in econometrics and are used in methodological contributions as well as in applications.

We consider the combinations of the generic approaches in Section 3.1 with Random Forest and Lasso.¹² Following Chernozhukov et al. (2018), we apply cross-fitting to all

¹⁰Previous works concerned with estimating IATEs by modifying tree-based methods are Su et al. (2009), Athey and Imbens (2016) and Wager and Athey (2018).

¹¹By 'local' we mean that each point of the conditional mean function is approximated by the (weighted) average of neighbouring observations. By 'global' we mean the attempt to approximate the conditional mean by a flexible functional fitted to all data simultaneously.

¹²We consider only 'pure' combinations where all estimation steps are conducted with one of the two machine learning methods and neglect the possibilities to estimate, e.g., the nuisance parameters via Random Forests and the IATEs via Lasso. In principle, this is possible but not pursued due to computational constraints. For the same reason, we do not pursue ensemble methods that combine

Table 2: List of considered causal machine learning estimators

	Random Forest	Lasso	Cross-fitting
Infeasible benchmark	X	X	
Conditional mean regression	X	X	
MOM IPW	X	X	X
MOM DR	X	X	X
MCM		X	X
MCM with EA		X	X
R-learning		X	X
Causal Forest	X		
Causal Forest with local centering	X		x

approaches that require the estimation of nuisance parameters. This means that the nuisance parameters and the IATEs are estimated in different samples to avoid overfitting.

Table 2 summarizes all estimators under investigation. We are currently not aware of a Random Forest implementation that supports weighted minimization. Thus, we implement MCM and R-learning only with Lasso. We add further an *infeasible benchmark* estimator that has access to the true ITEs and uses them as outcome in a standard prediction problem.

Like all machine learning methods, Random Forests and Lasso involve a variety of choices in the implementation. The following sections briefly explain Random Forests and Lasso, present the details of the implementation and explain the use of cross-fitting. The resulting estimators target the estimation of IATEs. Additionally, we consider the methods to estimate GATEs and ATE as a computationally cheap by-product of our analysis. To this end, we average the estimated IATEs, $\hat{\tau}(x)$, to GATEs by $\hat{\tau}(g) = N_g^{-1} \sum_{i=1}^{N} \mathbb{1}[G_i = g]\hat{\tau}(X_i)$, with $N_g = \mathbb{1}[G_i = g]$, and to the ATE by $\hat{\tau} = N^{-1} \sum_{i=1}^{N} \hat{\tau}(X_i)$.¹³

different estimators for the nuisance parameters or for IATE estimation (see, e.g., Rolling, Velez, & Yang, 2018; Schuler et al., 2018).

¹³Specialized estimators for GATE and ATE might outperform these aggregate estimators but this is beyond the scope of this paper (see, e.g., Lee et al., 2017; Chernozhukov et al., 2018).

4.1 Random Forest

The building block of Random Forests for conditional mean estimation are regression trees (Breiman et al., 1984). Regression trees recursively partition the sample along covariates to minimize MSE of the outcome. This leads to the tree structure and the means in the final leaves are used as predictions (for an introduction, see Hastie et al., 2009). However, regression trees are unstable and exhibit a high variance. The Random Forest of Breiman (2001) addresses this issue by combining a large number of decorrelated regression trees. The decorrelation is achieved by growing each tree on a random subsample (generated either by bootstrapping or subsampling) and by randomly choosing the covariates for each split decision.

The standard regression and probability forests split the trees to minimize the MSE of the observed outcomes. Those trees can then be used to form predictions for a realization of X_i . These predictions are formed as a weighted average of the observed outcomes where the weights are larger, the more often the observed outcome shares a final leave with the realization of X_i . These kind of forests are required for the conditional mean regression and the modified outcome approaches.

The Causal Forest of Athey et al. (2018) follows a similar structure. However, instead of splitting the sample according to observed outcomes, Causal Forests split the samples along the gradient of the mean difference with the pseudo outcomes

$$\rho_i = (D_i - \bar{D}_P)(Y_i - \bar{Y}_P - (D_i - \bar{D}_P)\hat{\beta}_P)/Var_P(D_i), \tag{17}$$

where \bar{D}_P and \bar{Y}_P are averages of treatment indicator and outcome, $\hat{\beta}_P$ is the mean difference and $Var_P(D_i)$ is the variance of the treatment in the parent node. This splitting rule is tailored to maximize heterogeneity and produces splits that are used to calculate the weights for the weighted mean comparison in Equation 16.

We also consider the Causal Forest with local centering. This means that D_i and Y_i in Equation 17 are replaced by $D_i - p(X_i)$ and $Y_i - \mu(X_i)$, respectively. The nuisance parameters are again estimated in a first step and this partialling out should remove the confounding at a high level before building the Causal Forest. Athey et al. (2018) show

that local centering can improve the performance of Causal Forests substantially in the presence of confounding.

We implement the regression forests for the conditional mean regression and the modified outcomes as well as the Causal Forests using the R package grf (Athey et al., 2018). We provide the forests with 105 baseline covariates for the prediction and set the number of variables that are considered at each split to 70. The minimum leaf size is set to one and one forests consists of 1,000 trees. We build honest trees such that the building of the tree and the estimation of the parameters are conducted in separate samples (Athey & Imbens, 2016). To this end, we split the sample randomly for each tree in three parts: 25% are used to build the tree, 25% to calculate the predictions, and 50% are left out.

4.2 Lasso

The Lasso is a shrinkage estimator and can be considered as an OLS estimator with a penalty on the sum of the absolute coefficients. The standard least squares Lasso solves the following minimization problem

$$\min_{\beta} \left[\sum_{i=1}^{N} w_i \left(Y_i - X_i \beta \right)^2 \right] + \lambda \sum_{j=1}^{p} \left| \beta_j \right|, \tag{18}$$

where w_i are weights, p is the number of covariates and λ is a tuning parameter to be optimally chosen.¹⁴ We obtain the standard OLS coefficients if the penalty term is equal to zero and we have at least as many observations as covariates. For a positive penalty term, at least some coefficients are shrunken towards zero to satisfy the constraint. The Lasso serves as a variable selector because some coefficients are set to zero if the penalty term is sufficiently increased. By incrementing the penalty term to a sufficiently large number, eventually all coefficients besides the constant are zero. The idea of this procedure is to shrink those variables with little or no predictive power to zero and to use the remaining shrunken coefficients for prediction. The degree of shrinkage should be chosen to balance the bias-variance trade-off and is the crucial tuning parameter of the Lasso.

We apply the R package glmnet to produce the predictions in the different approaches

¹⁴For the estimation of the propensity score, we use the equivalent logistic regression.

(Friedman, Hastie, & Tibshirani, 2010). We provide the estimator a set of 1,749 potential covariates including second order interactions and up-to fourth order polynomials.¹⁵ The tuning parameter is selected via 10-fold cross-validation.

4.3 Cross-fitting

Some approaches require the estimation of the nuisance parameters in a first step. We follow Chernozhukov et al. (2018) and apply cross-fitting to remove bias due to overfitting that is induced if nuisance and main parameters are estimated using the same observations. We implement a 50:50 version of their DML1 procedure. This means that we split the sample in two parts of the same size. In the first half, we estimate models for the nuisance parameters. We take these models to predict the nuisance parameters in the second half. These predicted nuisance parameters are then used in the estimation of the IATEs, $\hat{\tau}_1(x)$. We reverse the role of the two halves to obtain $\hat{\tau}_2(x)$. The estimates of the IATE are then calculated as $\hat{\tau}(x) = 1/2(\hat{\tau}_1(x) + \hat{\tau}_2(x))$.

4.4 Alternative estimation approaches

Table 2 above lists all estimators that we consider in this study. This list comprises not all alternatives, because we are not able to consider all estimators that have been proposed or would be possible combinations of the generic approaches and existing machine learning methods. We do not consider methods that are tailored for experimental studies (e.g., Imai & Ratkovic, 2013; Grimmer, Messing, & Westwood, 2017; McFowland, Somanchi, & Neill, 2018). Furthermore, restrictions in computation power force us to commit to approaches where we can leverage synergies in the implementation as illustrated in Figure C.1 of Appendix C. Thus, we are not able to consider the X-learner of Künzel et al. (2017), the three conditional outcome difference methods proposed by Powers et al. (2018), Orthogonal Random Forests (Oprescu, Syrgkanis, & Wu, 2018), Penalized Causal Forests (Lechner, 2018), methods based on neural nets (e.g., Johansson et al., 2016; Shalit et al., 2016; Schwab et al., 2018), Bayesian approaches like those based on Bayesian additive

¹⁵We exclude binary variables that represent less than 1% of the observations. Furthermore, we keep only one variable of variable combinations that show correlations of larger magnitude than ± 0.99 to speed up computation.

regression trees (BART) (Hill, 2011; Hahn, Murray, & Carvalho, 2017) or Bayesian forests (Taddy et al., 2016), and potentially other approaches that we are currently not aware of.

Further, the generic approaches discussed in Section 3.1 could be implemented using different machine learning algorithms like Boosting, Elastic Nets, Neural Nets, Ridge or any other supervised machine learning algorithm that minimizes the required loss functions (see for an overview Hastie et al., 2009).

5 Simulation set-up

5.1 Previous Empirical Monte Carlo Study

The simulation study of Wendling et al. (2018) is close in spirit to our approach, in the sense that their and our DGP relies as much as possible on real data. They compare eight conditional outcome difference estimators for binary outcomes, i.e., they focus on probability models. Their four DGPs are based on the covariates and the observational treatment assignment of four medical datasets. Thus, the IATEs and the resulting binary potential outcomes are the only components that need to be specified. The outcomes are simulated based on predictions of $\mu(0,x)$ and $\mu(1,x)$ from logistic neural networks (for more details, see Wendling et al., 2018). This is a realistic approach in the medical context. However, it removes two important features from the true outcome generating process. First, the projection of the outcome on observable covariates removes the impact of unobservable variables. Second, the true error structure is lost by imposing a logistic error term. Our EMCS aims to preserve these features of the data at least for the non-treatment outcome. Wendling et al. (2018) find that conditional mean regressions (implemented with BART, see Hill, 2011) and causal boosting (Powers et al., 2018) perform consistently well, while causal MARS (Powers et al., 2018) and Causal Forests (Athey et al., 2018) are found to be competitive for complex IATE but perform poorly if the variance of the IATE is relatively low. We implement conditional mean regressions and Causal Forests, but omit causal boosting and MARS because of computational restrictions.

¹⁶Nie and Wager (2017) use a similar EMCS for binary outcomes to assess the performance of different implementations of R-learning. However, they do not estimate the IATE but specify it to depend on two covariates.

5.2 Empirical Monte Carlo Study

Similar to Wendling et al. (2018) for the medical context, our study aims to approximate a real application in economic policy evaluation as close as possible. The idea of an EMCS is introduced by Huber et al. (2013) and Lechner and Wunsch (2013). It aims to take as many components of the DGP as possible from real data. We build this EMCS on 96,298 observations of Swiss administrative social security data that is already used in previous evaluation studies (Behncke, Frölich, & Lechner, 2010a, 2010b; Huber, Lechner, & Mellace, 2017). In particular, the EMCS mimics an evaluation of job search programs as in Knaus et al. (2017). The interest is in the heterogeneous effects of such a program on employment over the 33 months after the program start.¹⁷

Before we describe and motivate our EMCS approach, we list the general steps to evaluate estimators for IATEs, GATEs and ATEs in Table 3. We leave out a validation sample (10,000 randomly drawn observations) to compare the estimated IATEs against the true ITEs, while previous EMCS compare in-sample estimates to true values of a known IATE. This modification is intended to focus on the out-of-sample predictive power of the estimated causal effects. The advantage of this procedure is that we can specify the ITEs as ground truth without knowing the IATE as we describe below.

After removing the 10,000 observations of our validation sample, the remaining 78,844 observations form our 'population' from which we draw random subsamples of size 1,000 and 4,000 for estimation. We replicate this 2,000 times for the smaller and 500 times for the larger samples. The precision of the estimators and the computational costs increase with the sample size. Thus, we reduce the number of replications when we increase the sample size to restrict the latter. In case of \sqrt{N} -convergence, this will keep the simulation error approximately constant. Table 5 below shows the variants of the DGP for different N_s , R, p(x) and ξ_i . Before, we explain the specification of the two latter functions.

5.2.1 Propensity score

The 'population' propensity score is estimated in the full sample with 7,454 treated and 88,844 controls. After this estimation step, all treated are removed from the sample. The

 $[\]overline{^{17}}$ Appendix A provides more details about the outcomes and the rest of the dataset.

- 1. Take the full sample and estimate the propensity score, $p^{full}(x)$, using the method and specification of choice.
- 2. Remove all treated and keep only the N_{nt} non-treated observations. This means that Y_i^0 is observed for all members of the remaining subpopulation.
- 3. Specify the true ITEs, ξ_i .
- 4. Calculate the potential outcome under treatment as $\hat{Y}_i^1 = Y_i^0 + \xi_i$ for all observations.
- 5. Set aside a random validation sample of N_v observations. Remove this validation sample from the main sample.
- 6. Calculate any other parameters of interest in the validation sample as benchmark. Like GATEs as $\tau(g) = \left(\sum_{i=v}^{N_v} \mathbb{1}[G_v = g]\right)^{-1} \sum_{v=1}^{N_v} \mathbb{1}[G_v = g]\xi_v$ or ATEs as $\tau = N_v^{-1} \sum_{v=1}^{N_v} \xi_v$.
- 7. Draw a random sample of size N_s from the remaining $N_{nt} N_v$ observations.
- 8. Simulate pseudo treatment indicators $D_i \sim Bernoulli(p^{sim}(x))$, where $p^{sim}(x)$ is a potentially modified version of $p^{full}(x)$ to control the ratio of treated and controls or other features of the selection process.
- 9. Use the observation rule in 1 to create the observable outcome Y_i .
- 10. Use the N_s observations to estimate $\tau(x)$ with all estimators of interest.
- 11. Predict $\hat{\tau}(x)$ for all observations in the validation sample and use them to calculate $\hat{\tau}(q)$ as well as $\hat{\tau}$ for each estimator.
- 12. Repeat steps 7 to 11 R times.
- 13. Calculate performance measures.

specification of the propensity score is taken from Huber et al. (2017) and estimated using a standard logistic regression. We manipulate the constant to create a 50:50 split into treated and non-treated in the simulated samples. Appendix B.1 provides the details of the specification of the original propensity score and the distribution of the modified propensity score.

5.2.2 Specification of ITE

We are not able to observe the ITEs or any of its aggregates in a real world dataset. Therefore, we either need to estimate or to specify them. We choose the latter because estimation might favor similar estimators under investigation. Thus, our goal is to create

¹⁸We remove the 342 observations with a modified propensity score below 5% and above 95%. We deviate at this point from the real dataset and make the problem better behaved than in reality in terms of common support (see discussion in, e.g., Lechner & Strittmatter, 2017). We leave the investigation of performance in the presence of unbalanced ratios and insufficient common support for future studies and focus here on a relatively nice setting to start with.

a challenging synthetic ITE that uses components from real data. In observational studies, the estimators must be able to disentangle selection bias and effect heterogeneity. We make it hard for the estimators by using the 'population' propensity score $p^{HLM}(x)$ directly to calculate the ITEs. To this end, the propensity score is normalized and put into a sine function

$$\omega(x) = \sin\left(1.25\pi \frac{p^{HLM}(x)}{\max(p^{HLM}(x))}\right) + \varepsilon_i,\tag{19}$$

where ε_i is random noise. This highly non-linear function of the propensity score is standardized to have mean zero and variance one before it is scaled by the parameter α :

$$\Omega(x) = \alpha \frac{\omega(x) - \bar{\omega}}{SD(\omega(x))},\tag{20}$$

where $\bar{\omega}$ is the mean of $\omega(x)$ and $SD(\omega(x))$ is its standard deviation. Finally, we force the ITEs to respect two features of our outcome variable. This means that they are rounded to the next integer and that they must respect that \hat{Y}_i^1 falls between zero and 33.¹⁹ Thus, the final ITEs take the form

$$\xi(x, y^{0}) = \begin{cases} \lfloor \Omega(x) \rceil & \text{if } 0 \leq y^{0} + \lfloor \Omega(x) \rceil \leq 33 \\ -y^{0} & \text{if } y^{0} + \lfloor \Omega(x) \rceil < 0 \\ 33 - y^{0} & \text{if } y^{0} + \lfloor \Omega(x) \rceil > 33, \end{cases}$$
 (21)

where $\lfloor \cdot \rfloor$ indicates that we round to the nearest integer.

 $\xi(x,y^0)$ is highly non-linear and complicated due to the logistic function, the sine function and the rounding. Additionally, enforcement of the censoring makes it dependent on Y_i^0 that is taken directly from the data and thus depends on the covariates in an unknown fashion. Thus, we know the true ITEs but we do not know the functional form of the true IATE.

The true ITEs depend on the observables X_i and additionally on some unobservables

¹⁹The histogram of the (observed) Y_i^0 is provided in Figure A.1. Given the censored and integer nature of the outcome, we considered also using Poisson Lasso to estimate the outcome regressions. However, the computation time compared to least squared Lasso is substantially longer, while the predictive performance is very similar for our outcomes. Thus, we chose the least squares version for the simulations.

through $Y_i^{0.20}$ This means that the estimators approximate $\xi(x,y^0)$ using observables and produce estimates $\hat{\tau}(x)$ of $\tau(x)$. The goal of the EMCS is to figure out which estimators approximate the ITE comparatively well in this arguably realistic setting. The relative performance of the estimators translate then directly into the ability to approximate the unknown IATEs because estimators that minimize the MSE of the ITE also minimize the MSE of the IATE (see, e.g., Künzel et al., 2017). Note that the aggregation of IATEs to GATEs and ATE in step 6 of Table 3 can be considered as true values because the influence of Y_i^0 is averaged out for them asymptotically. This implies that the MSE of GATEs and ATEs would be approximately zero if $\hat{\tau}(x) = \tau(x)$, while the MSE of ITEs might still be positive in this case.

We consider three different values of α in Equation 31 to vary the size of the ITEs: $\alpha=0$ (ITE0), $\alpha=2$ (ITE1) and $\alpha=8$ (ITE2). Additionally, we create one specification without random noise and one with error term in Equation 30, $\varepsilon_i=0$ and $\varepsilon_i\sim 1-Poisson(1)$, respectively. Table 4 reports basic descriptive statistics of the resulting potential outcomes, ITEs, and GATEs. ITE0 without random noise creates a benchmark scenario that is most likely to be informative about which estimators are prone to confuse effect heterogeneity with selectivity. ITE1 leads to a scenario with moderate variance of the resulting ITEs. Their standard deviation amounts to about 14% of the non-treatment outcome. ITE2 produces bigger ITEs with a standard deviation of about 6, which is roughly 50% of the standard deviation of the non-treatment outcome. Thus, they should be less difficult to detect.

The ITEs without and with random noise are created to be similar in their first two moments. However, they differ substantially in their variation that can be explained by observables. The influence of Y_i^0 on the ITEs is substantial because between 27% and 44% of the observations are censored for the non-zero ITEs. This explains why ITE1 and ITE2 without random noise are not deterministic either and therefore not perfectly predictable by X_i . Still, the out-of-sample R^2 of Random Forest and Lasso predictive regressions shown in Table B.2 document that we can explain between about 50% and 70% of the

 $[\]overline{^{20}}$ These unobservables do not invalidate the CIA in our simulated samples, as Y_i^0 and thus the unobservables are not part of the population propensity score. The alternative to ensure a valid CIA in an EMCS is to keep the true treatment allocation structure and to specify the potential outcomes as function of the observables. This is the approach of Wendling et al. (2018) that is discussed in 5.1.

Table 4: Descriptive statistics of simulated outcomes and ITEs

	Mean	Std. Dev.	Skewness	Kurtosis	Percent censored			
Without random	Without random noise $(\varepsilon_i = 0)$:							
Y^0 in all DGPs			-0.1	1.4	-			
Y^1 in ITE0	16.1	12.8	-0.1	1.4	-			
Y^1 in ITE1	16.3	12.6	-0.1	1.4	-			
Y^1 in ITE2	16.3	12.6	0.1	1.5	-			
ITE0	0.0	0.0	-	-	0.0			
ITE1	0.1	1.8	-0.3	2.3	39.2			
ITE2	0.2	6.4	-0.4	2.5	43.7			
GATE0	0.0	0.0	-	-	-			
GATE1	-0.4	1.8	-0.1	2.0	-			
GATE2	-1.8	6.2	-0.3	2.1	-			
With random no	$vise (\varepsilon_i \sim$	$\sim 1 - Poisso$	on(1)):					
Y^1 in ITE0	16.2	12.7	-0.1	1.4	-			
Y^1 in ITE1	16.3	12.6	-0.1	1.4	-			
Y^1 in ITE2	16.5	12.3	0.0	1.5	-			
ITE0	0.1	0.9	-1.2	5.1	26.6			
ITE1	0.1	1.8	-1.0	4.9	36.7			
ITE2	0.3	6.3	-1.0	4.8	41.1			
GATE0	0.0	1.1	-1.2	3.5	-			
GATE1	0.0	1.7	-0.7	3.2	-			
GATE2	0.0	5.8	-0.6	3.7	-			

Notes: Potential outcomes and ITEs are considered for all observations. GATEs are considered for the validation sample.

ITEs with our covariates. With random noise, this explainable part decreases to close to zero for ITE0 and 6.3% for ITE1. We consider the latter to be a more realistic scenario because the individual component is expected to be relatively large.²¹ Thus, we select ITE1 and ITE2 with random noise as our baseline DGPs additionally to the benchmark scenario ITE0 without random noise.

The first column of Table 4 shows that we specify the mean of the ITEs, the ATE, close to zero.²² Appendix B.3 describes how we aggregate the ITEs into 64 groups with sizes between 32 and 420 observations to specify the true GATEs.

In summary, we consider six different scenarios defined by different choices for the scale

²¹For example, Djebbari and Smith (2008) provide evidence that the ITEs in their applications show only little systematic variation.

²²Appendix B.2 shows in detail how the ITEs and potential outcomes are distributed, how the ITEs are related to the propensity score and Y_i^0 , as well as an interpretation of the simulated selection behavior of caseworkers. Note that the lower standard deviations of Y_i^1 compared to Y_i^0 result from the censoring that moves mass away from the bounds (see Figures B.4 and B.5).

Table 5: List of DGPs

	N_s	α in 31	Propensity score	R	ε_i in 30				
With selection	With selection and without random noise:								
ITE0*	1,000	$\alpha = 0$	$p^{HLM}(x)$	2,000	$\varepsilon_i = 0$				
ITE1	1,000	$\alpha = 2$	$p^{HLM}(x)$	2,000	$\varepsilon_i = 0$				
ITE2	1,000	$\alpha = 8$	$p^{HLM}(x)$	2,000	$\varepsilon_i = 0$				
ITE0*	4,000	$\alpha = 0$	$p^{HLM}(x)$	500	$\varepsilon_i = 0$				
ITE1	4,000	$\alpha = 2$	$p^{HLM}(x)$	500	$\varepsilon_i = 0$				
ITE2	4,000	$\alpha = 8$	$p^{HLM}(x)$	500	$\varepsilon_i = 0$				
With selection	on and rai	ndom nois							
ITE0	1,000	$\alpha = 0$	$p^{HLM}(x)$	2,000	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE1*	1,000	$\alpha = 2$	$p^{HLM}(x)$	2,000	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE2*	1,000	$\alpha = 8$	$p^{HLM}(x)$	2,000	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE0	4,000	$\alpha = 0$	$p^{HLM}(x)$	500	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE1*	4,000	$\alpha = 2$	$p^{HLM}(x)$	500	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE2*	4,000	$\alpha = 8$	$p^{HLM}(x)$	500	$\varepsilon_i \sim 1 - Poisson(1)$				
With random assignment and without random noise:									
ITE0	1,000	$\alpha = 0$	0.5	2,000	$\varepsilon_i = 0$				
ITE1	1,000	$\alpha = 2$	0.5	2,000	$\varepsilon_i = 0$				
ITE2	1,000	$\alpha = 8$	0.5	2,000	$\varepsilon_i = 0$				
ITE0	4,000	$\alpha = 0$	0.5	500	$\varepsilon_i = 0$				
ITE1	4,000	$\alpha = 2$	0.5	500	$\varepsilon_i = 0$				
ITE2	4,000	$\alpha = 8$	0.5	500	$\varepsilon_i = 0$				
With randon	n $assignm$	ent and r	andom noise:						
ITE0	1,000	$\alpha = 0$	0.5	2,000	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE1	1,000	$\alpha = 2$	0.5	2,000	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE2	1,000	$\alpha = 8$	0.5	2,000	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE0	4,000	$\alpha = 0$	0.5	500	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE1	4,000	$\alpha = 2$	0.5	500	$\varepsilon_i \sim 1 - Poisson(1)$				
ITE2	4,000	$\alpha = 8$	0.5	500	$\varepsilon_i \sim 1 - Poisson(1)$				

Notes: Asteriks mark the baseline DGPs.

of the ITEs and the random noise variables. Additionally to the DGP with selection into the treatment, we consider also the case of an experiment with 50:50 random assignment. These twelve different DGPs are considered for the sample sizes of 1,000 and 4,000 observations leading to a total number of 24 different settings. Table 5 summarizes all parameter settings in which the eleven estimators are compared.

5.2.3 Performance measures

We consider three major performance measures: mean squared error (MSE), absolute bias (|Bias|) and standard deviation (SD) for the prediction of each observation v in the

validation sample:²³

$$MSE_v = \frac{1}{R} \sum_{r=1}^{R} \left[\xi(x_v, y_v^0) - \hat{\tau}(x_v)_r \right]^2$$
 (22)

$$|Bias_v| = |\underbrace{\frac{1}{R} \sum_{r=1}^{R} \hat{\tau}(x_v)_r}_{\bar{\hat{\tau}}(x_v)_r} - \xi(x_v, y_v^0)|$$
 (23)

$$SD_{v} = \sqrt{\frac{1}{R} \sum_{r=1}^{R} \left[\hat{\tau}(x_{v})_{r} - \bar{\hat{\tau}}(x_{v})_{r} \right]^{2}}$$
 (24)

Most simulation studies are interested in only few parameters such that the performance measure for each parameter can be reported and interpreted. However, in our case we have 10,000 parameters such that we need to summarize the performance over the whole validation sample by taking the averages \overline{MSE} , $|\overline{Bias}|$ and \overline{SD} . Additionally, we apply the Jarque-Bera test (JB) to the distribution of predictions for each observation v in the validation sample and report the fraction of observations for which normality is rejected at the 5% confidence level.²⁵

Results 6

6.1 IATE estimation

Table 6 shows the main performance measures for the three baseline DGPs.²⁶ At first, we compare estimators within similar approaches to identify the competitive versions. Afterwards, we compare the competitive versions over all approaches to identify those estimators that show an overall good performance and provide a general comparison of Random Forest and Lasso based methods.

 $^{^{23}}$ The formulas are written for the ITE. The same measures are used for GATE and ATE.

 $^{^{24}}$ For example, $\overline{MSE}=N_v^{-1}\sum_{v=1}^{N_v}MSE_v.$ 25 Appendix D discusses and provides also alternative performance measures.

²⁶The full tables with more performance measures are provided in Tables D.1 for ITE0, D.5 for ITE1, and D.6 for ITE2.

Table 6: Simulation results of ITE estimation for baseline DGPs

	1000 observations				4000 observations			
	\overline{MSE}	$ \overline{Bias} $	\overline{SD}	JB	\overline{MSE}	$ \overline{Bias} $	\overline{SD}	JB
	I'.	ΓE0 wit	h selec	ction an	d without	rando	m noise	е
Random Forest:								
Infeasible					ependent v			
Conditional mean regression	3.69	0.62	1.78	6%	2.79	1.33	1.49	4%
MOM IPW	10.52	2.05	2.16	18%	5.76	1.89	1.74	12%
MOM DR	2.00	0.40	1.35	7%	1.16	0.85	1.03	8%
Causal Forest	3.52	0.75	1.69	12%	2.31	1.21	1.29	6%
Causal Forest with local centering	3.42	0.34	1.81	10%	2.05	1.13	1.40	9%
Lasso: Infeasible		Ŋ	Jo veri	otion in d	ependent v	zaria blo		
Conditional mean regression	11.21	0.69	3.19	91%	6.14	1.92	2.30	35%
MOM IPW	11.21 11.31	1.09	$\frac{3.19}{2.99}$	100%	5.02	1.52 1.56	$\frac{2.30}{1.93}$	100%
MOM II W MOM DR	45.39	0.60	6.31	100%	0.51	0.51	0.62	35%
MCM MCM	13.03	1.50	3.05	100%	5.79	1.65	1.94	100%
MCM with efficiency augmentation	$\frac{13.03}{2.08}$	0.42	1.36	100% $100%$	$\begin{array}{c} 0.19 \\ 0.48 \end{array}$	0.49	0.62	97%
R-learning	2.03	$0.42 \\ 0.45$	1.30 1.33	100% $100%$	$\begin{array}{c} 0.48 \\ 0.47 \end{array}$	0.49 0.49	0.62	97%
n-learning	2.03							9170
		ITE1	with	selection	n and ran	dom no	oise	
Random Forest:	2.00	1.00	0.15	7107	0.02	1.20	0.11	2607
Infeasible	2.98	1.29	0.15	71%	2.93	1.30	0.11	26%
Conditional mean regression	7.04	1.45	1.78	8%	6.05	1.44	1.49	4%
MOM IPW	12.92	2.26	2.20	16%	8.42	1.76	1.77	11%
MOM DR	5.08	1.36	1.33	8%	4.17	1.32	1.01	9%
Causal Forest	6.86	1.49	1.68	12%	5.61	1.48	1.29	6%
Causal Forest with local centering Lasso:	6.50	1.35	1.79	12%	5.10	1.32	1.39	10%
Infeasible	3.00	1.28	0.21	100%	2.93	1.28	0.16	83%
Conditional mean regression	14.26	1.46	3.16	90%	9.20	1.43	2.30	36%
MOM IPW	15.69	1.56	3.12	100%	8.03	1.46	2.01	100%
MOM DR	48.76	1.40	6.32	100%	3.66	1.34	0.64	96%
MCM	15.31	1.72	3.10	100%	8.14	1.51	1.96	100%
MCM with efficiency augmentation	5.27	1.37	1.36	100%	3.62	1.33	0.63	98%
R-learning	5.16	1.38	1.29	100%	3.65	1.34	0.63	98%
		ITE2	with	selection	n and ran	dom no	oise	
Random Forest:								
Infeasible	38.46	4.43	0.52	66%	37.86	4.43	0.41	31%
Conditional mean regression	43.74	4.58	1.74	8%	42.26	4.54	1.46	5%
MOM IPW	46.83	4.83	2.23	16%	42.69	4.54	1.80	11%
MOM DR	41.45	4.50	1.32	10%	40.03	4.45	1.03	11%
Causal Forest	43.87	4.61	1.66	12%	42.34	4.58	1.29	7%
Causal Forest with local centering Lasso:	42.84	4.50	1.78	12%	41.05	4.46	1.40	9%
Infeasible	38.66	4.42	0.71	100%	37.84	4.40	0.53	84%
Conditional mean regression	50.11	$\frac{4.42}{4.52}$	3.15	92%	44.31	4.40 4.46	$\frac{0.55}{2.33}$	34%
MOM IPW	49.82	$\frac{4.52}{4.50}$	3.10	$\frac{92\%}{100\%}$	44.31 43.21	4.40 4.43	$\frac{2.33}{2.17}$	97%
MOM IF W MOM DR	49.82 537.16	$\frac{4.50}{4.55}$	5.20	100% $100%$	40.21 40.11	4.43	0.76	97%
MCM MCM	49.25	$\frac{4.55}{4.47}$	3.18	100% $100%$	40.11 42.63	4.46	$\frac{0.76}{2.07}$	100%
MCM with efficiency augmentation	49.25 41.99	$\frac{4.47}{4.51}$	$\frac{3.18}{1.41}$	100% $100%$	42.03 40.04	$\frac{4.41}{4.47}$	$\frac{2.07}{0.75}$	99%
R-learning	41.99 42.13	$\frac{4.51}{4.54}$	1.41 1.35	100% $100%$	40.04	$\frac{4.47}{4.49}$	$0.75 \\ 0.74$	99%
10-10ariiiig	42.10	4.04	1.00	100/0	40.20	4.49	0.74	<i>3</i> 0/0

Notes: \overline{MSE} shows the mean MSE of all 10,000 observations in the validation sample, $|\overline{Bias}|$ denotes the mean absolute bias, \overline{SD} the mean standard deviation, and JB the fraction of observations for which the Jarque-Bera test is rejected at the 5% level. Bold numbers indicate the best performing estimators in terms of \overline{MSE} and estimators within two standard (simulation) errors of the lowest \overline{MSE} .

6.1.1 Conditional mean regressions

The Random Forest version of conditional mean regressions clearly outperforms the Lasso version in terms of mean MSE. The differences are particularly striking in the smaller sample estimation of ITEO where the mean MSE of the Lasso version is more than three times larger compared to the Random Forest version. The substantially worse performance of the Lasso version is consistently observed over all baseline DGPs and sample sizes. This is mostly driven by a substantially lower mean SD of Random Forest based conditional mean regressions that is thus the dominant choice within the two considered versions of conditional mean regressions.

6.1.2 Modified outcome methods

The ranking of the MOM estimators depends on the sample size. Table 6 shows that Random Forests are superior to the Lasso versions in the smaller samples. Especially the DR modification with Random Forest performs well due to relatively low mean SD. In contrast, the Lasso equivalent is by far the worst estimator in the smaller samples. It has up to twice as large mean SD compared to the next worst estimator and shows consequently a very high mean MSE. One potential reason is that only the DR estimators require the estimation of $\mu(d, x)$ as a nuisance parameter. These predictions are then based on only 250 observations when using cross-fitting, while $\mu(x)$ and p(x) are based on 500 observations. The instability of the Lasso as outcome predictor in small samples seems to spillover to the IATE estimation.²⁷ The results for the larger sample size indicate that the poor performance is a small sample issue. The DR modification with Lasso outperforms the other versions of MOM in ITE0 and ITE1 and is also close to its Random Forest equivalent for ITE2. The good performance is mainly driven by relatively low mean SDs.

As expected from the results of Powers et al. (2018), the IPW modification has relatively high mean SD and is therefore not competitive. This is despite the fact that our DGP does not lead to extreme propensity scores and creates thus a relatively favorable setting

²⁷Chernozhukov et al. (2018) observe a similar problem of global approximations for the estimation of average effects. See also Waernbaum and Pazzagli (2017) for conditions under which a poor approximation of the outcome leads to worse performance of DR estimators compared to IPW. Similarly, Kang and Schafer (2007) demonstrate that double robust methods can perform poorly if both nuisance parameters are misspecified.

for IPW. Therefore, the DR modification seems to be in general the dominant choice as long as the Lasso version is not used in small samples.

6.1.3 MCM and R-learning

The results of the three estimators with modified covariates are similar to the results for the MOM. The MCM is clearly outperformed by its efficiency augmented version and R-learning that both use the outcome regression additionally to the propensity score as a nuisance parameter. For MCM, the efficiency augmentation more than halves the mean SD in all baseline DGPs. Thus, the additional computational effort is fruitful when using MCM. For all DGPs, efficiency augmented MCM and R-learning perform very similar along all dimensions. This finding is in line with the synthetic simulation in Appendix D of Chen et al. (2017) who find also very similar results for those two.

6.1.4 Causal Forests

The Causal Forest is specialized to maximize heterogeneity in experimental settings but it is not build to explicitly account for selection. Thus, it is prone to choose splits that do not sufficiently remove selection bias. However, Causal Forests with local centering address this problem by partialling out the selection effects in a first step. They are specialized to maximize effect heterogeneity and to account for selection bias. Consequently, they uniformly perform better than Causal Forests. This is driven by a relatively low mean absolute bias, but higher mean SD partly offsets this advantage. However, the differences between the two Causal Forest versions are moderate but the version with local centering is the dominant choice if the goal is to minimize mean MSE. However, the improvement comes at the cost of estimating additionally two nuisance parameters before estimating the Causal Forest.²⁸

²⁸Together with the conditional mean regression based on Random Forests, the Causal Forest is thus attractive if computation time is a concern. Appendix D.4 shows that both require very similar computation time and are the fastest Random Forest based estimators under consideration.

6.1.5 Overall comparison

The results in Table 6 show that no estimator is uniformly superior for all sample sizes and DGPs. However, we can categorize the estimators into those that show a relatively good performance in all settings, the volatile ones with outstanding performance only in particular settings, and those that are never competitive.

The first category comprises Random Forest MOM DR, MCM with efficiency augmentation, R-learning and Causal Forest with local centering. These four estimators are in a similar range over all DGPs and sample sizes and belong consistently to the five best estimators. Thus, they seem to be reasonable choices to estimate IATEs. Causal Forest with local centering is the only one of those four that shows never the best performance in terms of mean MSE. This is driven by a larger mean SD that works against the very competitive mean absolute bias. The feature that unifies all four best performing estimators is that they use propensity score and outcome regressions as nuisance parameters in the estimation process.

The MOM DR with Lasso belongs to the second category because it is very competitive for larger samples but the worst choice in smaller samples. Thus, it remains a risky choice for applications because the critical sample size for good performance may depend on the particular dataset.

Finally, conditional mean regressions, MOM IPW with Lasso and Causal Forest should not be considered in settings like ours if minimizing MSE has a high priority. However, if computational constraints are binding, conditional mean regressions with Random Forests and Causal Forests can be attractive options.

6.1.6 Random Forest vs. Lasso

A direct comparison of Random Forest and Lasso is possible for conditional mean regressions and MOM. For the smaller sample size, Random Forest clearly outperforms the Lasso based versions. This is driven by the substantially lower mean SD of Random Forest based estimators. The reason is that the global approximations of Lasso are rather instable for small samples. This instability is reduced for larger samples and the Lasso based MOM performs better than the Random Forest equivalents for ITE0 and ITE1.

This dependence on the sample size is not observed for the Lasso specific estimators MCM with efficiency augmentation and R-learning. Both show competitive performance regardless of the sample size. This is particularly surprising given the highly non-linear ITEs. However, all Lasso based methods are far from being normally distributed. For at least 30% of the validation observations, the JB test rejects normality. For many estimators it is even rejected for all observations in the validation sample, while we would expect only a fraction of 5% to be rejected under normality. Columns 9 in the Tables of Appendix D show that this is due to excess kurtosis, which indicates that the Lasso based methods are prone to produce outliers. It is mitigated for the sample size of 4,000 but still the JB test is rejected for a large majority. This reflects the theoretical results of Leeb and Pötscher (2005, 2008) that shrinkage estimators like Lasso exhibit non-normal finite sample distributions.

In contrast, all Random Forest based estimators appear to be approximately normally distributed. This is also reflected by a mean skewness close to zero and a kurtosis close to three. Decently performing Random Forest based estimators might be therefore preferable to slightly better performing Lasso based estimators. The former produce less outliers and seem therefore more reliable and robust in empirical applications as well as more amendable to statistical inference.

6.2 GATE and ATE estimation

Table 7 shows the main performance measures of GATE estimation for the three baseline DGPs.²⁹ We observe similar patterns as for the IATE estimation and the categorization of estimators in Section 6.1.5 remains by and large the same. The four estimators that show a consistently good performance for IATEs are also good choices for the estimation of GATEs.

For GATE estimation, we observe a new candidate with outstanding performance in a particular setting. MCM performs remarkably well for ITE2 showing the second lowest mean MSE. The mean absolute bias of MCM is already competitive for the estimation of

²⁹The full tables with more performance measures are provided in Tables D.15 for ITE0, D.19 for ITE1, and D.20 for ITE2. The results for all DGPs are provided in Appendix D.2.

Table 7: Simulation results of GATE estimation for baseline DGPs

	1000 observations				4000 observations			
	\overline{MSE}	$ \overline{Bias} $	\overline{SD}	JB	\overline{MSE}	$ \overline{Bias} $	\overline{SD}	JB
GA	TEs from	m ITE0	with	selection	and wit	hout ra	ndom	noise
Random Forest:								
Conditional mean regression	1.77	0.55	1.19	22%	1.07	0.49	0.86	8%
MOM IPW	4.44	1.59	1.16	47%	1.12	0.67	0.66	6%
MOM DR	0.87	0.38	0.85	20%	0.30	0.25	0.48	14%
Causal Forest	1.44	0.70	0.96	17%	0.74	0.64	0.53	0%
Causal Forest with local centering	1.08	0.33	0.99	8%	0.35	0.22	0.54	3%
Lasso:								
Conditional mean regression	3.35	0.55	1.70	34%	1.45	0.47	1.06	6%
MOM IPW	3.15	0.78	1.50	100%	1.19	0.53	0.87	86%
MOM DR	38.85	0.59	6.20	100%	0.30	0.31	0.45	38%
MCM	4.56	1.20	1.62	100%	1.65	0.75	0.92	94%
MCM with efficiency augmentation	1.04	0.41	0.93	66%	0.27	0.26	0.45	55%
R-learning	1.04	0.44	0.92	73%	0.27	0.28	0.44	25%
					election a			
D 1 E 1		125 1101		I WIUII BO		na ranc		31 5C
Random Forest:	0.20	0.05	1 10	2007	1 59	0.76	0.00	6%
Conditional mean regression	2.30	0.85	1.18	20%	1.53	0.76	0.86	
MOM IPW	3.84	1.41	1.17	41%	0.99	0.54	0.67	11%
MOM DR	1.16	0.59	0.83	20%	0.49	0.42	0.48	17%
Causal Forest	2.04	1.01	0.95	19%	1.26	0.93	0.53	2%
Causal Forest with local centering	1.38	0.56	0.97	17%	0.58	0.44	0.54	5%
Lasso:				~				
Conditional mean regression	3.68	0.78	1.69	41%	1.66	0.60	1.06	14%
MOM IPW	3.03	0.65	1.53	100%	1.17	0.47	0.89	77%
MOM DR	39.33	0.79	6.20	100%	0.59	0.50	0.46	42%
MCM	3.98	0.93	1.65	97%	1.31	0.52	0.93	91%
MCM with efficiency augmentation	1.40	0.62	0.93	80%	0.55	0.47	0.45	39%
R-learning	1.45	0.68	0.91	75%	0.61	0.51	0.45	48%
	GAT	Es fron	n ITE	with selection and random noise				
Random Forest:								
Conditional mean regression	5.28	1.57	1.15	20%	3.97	1.44	0.85	22%
MOM IPW	3.75	1.25	1.18	41%	1.72	0.95	0.68	14%
MOM DR	3.50	1.29	0.84	23%	2.19	1.08	0.49	23%
Causal Forest	5.33	1.71	0.94	25%	4.17	1.60	0.53	8%
Causal Forest with local centering	3.74	1.28	0.98	11%	2.43	1.12	0.55	9%
Lasso:								
Conditional mean regression	5.73	1.34	1.71	42%	2.59	0.95	1.09	11%
MOM IPW	4.00	1.01	1.59	100%	1.92	0.82	0.96	45%
MOM DR	30.36	1.43	4.71	100%	2.71	1.23	0.52	69%
MCM	3.65	0.66	1.67	100%	1.58	0.61	0.97	81%
MCM with efficiency augmentation	3.94	1.35	0.95	72%	2.63	1.22	0.51	67%
v o								50%
R-learning	3.94 4.27	1.35	0.95	72% 72%	2.63 2.95	1.22	$0.51 \\ 0.50$	

Notes: \overline{MSE} shows the mean MSE of all 10,000 observations in the validation sample, $|\overline{Bias}|$ denotes the mean absolute bias, \overline{SD} the mean standard deviation, and JB the fraction of observations for which the Jarque-Bera test is rejected at the 5% level. Bold numbers indicate the best performing estimators in terms of \overline{MSE} and estimators within two standard (simulation) errors of the lowest \overline{MSE} .

IATEs of ITE2 in Table 6. However, the mean SD is more than twice as large compared to the best estimators, which prevents a competitive performance in terms of mean MSE. The averaging of these noisy but relatively unbiased estimators seems to produce a competitive estimator for the higher aggregation level. Still, MCM performs poorly for the other DGPs.

The averaging improves also the performance of locally centered Causal Forests relative to its uncentered version. The results for the estimation of IATEs show that the advantage in terms of mean absolute bias is partly offset by a higher variability. The aggregation step reduces this difference such that the lower bias translates also into a substantially lower mean MSE.

The aggregation leads further to a substantial reduction in the excess kurtosis of all Lasso estimators (see Tables of Appendix D.2). However, the JB test is still rejected for most observations. Note that we observe for all estimators a substantial amount of bias although the influence of Y_i^0 and the irreducible noise is averaged out to a large extent. This indicates that the estimators are not able to completely remove the selection bias, which is particularly problematic if we would be interested in statistical inference. The results for ITE0 without noise and with random assignment in Appendix D.2 provide evidence in this direction.

The results for the estimation of ATEs in Appendix D.3 are mostly in line with those for GATEs. Again, MCM is highly competitive and provides the best performing estimators for ITE2. Also the benefits of averaging the locally centered Causal Forest are observed. The bias is halved compared to their uncentered version while both show similar SDs.

The skewness and kurtosis show that the ATE estimators are mostly normally distributed with mean skewness close to zero and mean kurtosis close to three. The obvious exception is MOM DR with Lasso for which also averaging the IATEs does not mitigate the bad performance due to extreme outliers.

Finally, we note that the comparison of the mean MSE for the two sample sizes indicates that GATEs and ATEs estimators show a substantially faster convergence rate compared to the respective IATE estimators. This indicates that the additional averaging of noisily estimated IATEs results in faster convergence and the ATEs may be estimable with close to parametric rates. However, we do not overemphasize this finding as it is only based on

two sample sizes.

6.3 Alternative DGPs

The discussions in the previous section focus on the results of the three baseline DGPs. This section summarizes the major insights from the alternative DGPs. Their results are provided in Appendices D.1.1 to D.3.4 where we also discuss details and peculiarities of the specific DGPs and aggregation levels. In general, the four estimators that are identified as the best performing for the baseline DGPs belong also to the best performing ones in the alternative DGPs.

For the estimation of *IATEs*, we observe new candidates that are only successful in particular DGPs with selection into treatment. For example, Random Forest MOM IPW performs very good for ITE2 without noise. However, these and other peculiarities discussed in the Appendix hold only for either mean MSE or median MSE, while the four best performing candidates are usually competitive in both measures. Additionally, we assess whether our findings stay robust if we ignore the natural bounds of our outcome variable when creating the DGP. The results in Appendix D.1.5 show that this is the case in a DGP that allows treated outcomes outside the natural bounds of the original outcome when defining the ITEs.

We also consider all previously discussed DGPs with random treatment assignment. This means that the estimation problem becomes easier because selection bias is no longer a concern. By and large the results are in line with the respective results for the DGPs with selection into treatment, especially the conclusions about the best performing estimators are not changed. As expected, the mean MSEs for the DGP with random assignment are lower for most of the estimators and thus closer to the infeasible benchmark. This is always driven by a lower mean absolute bias while the mean SDs are very similar to the equivalent DGPs with selection. This suggests that the methods are not able to completely remove the selection bias.

Two other differences to the baseline DGPs are noteworthy. First, MOM DR with Lasso shows competitive performance already in small samples. This indicates that the bad performance is related to large errors made in the outcome *and* the selection equation

in small samples, which is in line with the simulation evidence of Kang and Schafer (2007) for DR ATE estimators. Second, Causal Forest and its locally centered version show a nearly identical performance. This illustrates that local centering is only beneficial when there is selection into treatment.

The results for *GATE* and *ATE* estimators confirm the observation in the baseline DGPs that IATE estimators with low bias but high variance can provide competitive estimators if they are averaged to higher aggregation levels. In particular, averaging MCM IATEs shows in many alternative DGPs a relatively good performance. However, MCM performs worst in some other DGPs in a non-systematic way. Thus, the results show that noise can be averaged out for these higher aggregation levels but there is no guarantee for this. The estimators that are already successful for the IATEs remain the most reliable choices for GATEs and ATEs.

In general, we find that the differences between the estimators become smaller, the more the IATEs are aggregated. Especially, the SDs become more similar by averaging IATEs such that the differences between the estimators are mainly driven by bias.

7 Conclusion

This is the first comprehensive simulation study in economics that investigates the finite sample performance of a large number of different causal machine learning estimators. We rely on arguably realistic DGPs that have potentially more external validity than the mostly synthetic DGPs considered so far in the limited simulation literature for these estimators. We consider DGPs with and without selection into treatment. Our main goal is to estimate individualized average treatment effects. Additionally, we report the performance of estimators that aggregate individualized average treatment effects to an intermediate and the population level.

We do not find any single causal machine learning estimator that consistently outperforms all other estimators. However, we do find a group of four estimators that show competitive performance for all DGPs. This group includes the Causal Forest with local centering, Random Forest based MOM DR, MCM with efficiency augmentation, and R-learning. These estimators explicitly use both, the outcome and the treatment equations in a multiple step procedure. The estimators that use the Lasso have heavy tails in the smaller samples.

The best performing estimators for the individualized average treatment effects produce also the most reliable estimators for higher aggregation levels. However, in some settings also noisily estimated individualized average treatment effects with low bias produce competitive estimators for higher aggregates because the noise is averaged out.

Despite relying as much as possible on arguably realistic DGPs, the external validity of every simulation study is uncertain. Future research will show if our findings hold in other empirical settings. Furthermore, it may be possible to improve the performance of each method with more tailored implementations. Finally, we focus in this study on the finite sample performance of point estimates and leave the investigation of inference procedures to future research.

References

- Abadie, A. (2005). Semiparametric difference-in-differences estimators. Review of Economic Studies, 72(1), 1–19. doi: 10.1111/0034-6527.00321
- Abrevaya, J., Hsu, Y.-C., & Lieli, R. P. (2015). Estimating conditional average treatment effects. *Journal of Business & Economic Statistics*, 33(4), 485–505. doi: 975555
- Andini, M., Ciani, E., de Blasio, G., D'Ignazio, A., & Salvestrini, V. (2018). Targeting with machine learning: An application to a tax rebate program in Italy. *Journal of Economic Behavior & Organization*, 156, 86–102.
- Ascarza, E. (2018). Retention futility: Targeting high risk customers might be ineffective.

 Journal of Marketing Research, 55(1), 80–98.
- Athey, S. (2018). The impact of machine learning on economics. In A. K. Agrawal, J. Gans, & A. Goldfarb (Eds.), *The economics of artificial intelligence: An agenda*. University of Chicago Press.
- Athey, S., & Imbens, G. W. (2016). Recursive partitioning for heterogeneous causal effects.

 Proceedings of the National Academy of Sciences, 113(27), 7353–7360.
- Athey, S., & Imbens, G. W. (2017). The state of applied econometrics: causality and policy evaluation. *Journal of Economic Perspectives*, 31(2), 3–32.
- Athey, S., Tibshirani, J., & Wager, S. (2018). Generalized random forests. *Annals of Statistics*, forthcoming.
- Behncke, S., Frölich, M., & Lechner, M. (2010a). A caseworker like me Does the similarity between the unemployed and their caseworkers increase job placements? The Economic Journal, 120(549), 1430–1459.
- Behncke, S., Frölich, M., & Lechner, M. (2010b). Unemployed and their caseworkers: Should they be friends or foes? *Journal of the Royal Statistical Society: Series A* (Statistics in Society), 173(1), 67–92.
- Bell, S. H., & Orr, L. L. (2002). Screening (and creaming?) applicants to job training programs: The AFDC homemaker home health aide demonstrations. *Labour Economics*, 9(2), 279–301.
- Belloni, A., Chernozhukov, V., Fernández-Val, I., & Hansen, C. (2017). Program evaluation

- and causal inference with high-dimensional data. Econometrica, 85(1), 233–298.
- Belloni, A., Chernozhukov, V., & Hansen, C. (2014a). High-dimensional methods and inference on structural and treatment effects. *Journal of Economic Perspectives*, 28(2), 29–50.
- Belloni, A., Chernozhukov, V., & Hansen, C. (2014b). Inference on treatment effects after selection among high-dimensional controls. *Review of Economic Studies*, 81(2), 608–650.
- Bertrand, M., Crépon, B., Marguerie, A., & Premand, P. (2017). Contemporaneous and post-program impacts of a public works program: Evidence from Côte d'Ivoire. Washington D.C.
- Breiman, L. (2001). Random forests. *Machine Learning*, 45(1), 5–32.
- Breiman, L., Friedman, J., Stone, C. J., & Olshen, R. A. (1984). Classification and regression trees. CRC press.
- Card, D., Kluve, J., & Weber, A. (2017). What works? A meta analysis of recent active labor market program evaluations. *Journal of the European Economic Association*, 16(3), 894–931.
- Chen, S., Tian, L., Cai, T., & Yu, M. (2017). A general statistical framework for subgroup identification and comparative treatment scoring. *Biometrics*, 73(4), 1199–1209.
- Chernozhukov, V., Chetverikov, D., Demirer, M., Duflo, E., Hansen, C., & Newey, W. (2017). Double/Debiased/Neyman machine learning of treatment effects. *American Economic Review Papers and Proceedings*, 107(5), 261–265.
- Chernozhukov, V., Chetverikov, D., Demirer, M., Duflo, E., Hansen, C., Newey, W., & Robins, J. (2018). Double/Debiased machine learning for treatment and structural parameters. *The Econometrics Journal*, 21(1), C1-C68.
- Davis, J. M., & Heller, S. B. (2017). Using causal forests to predict treatment heterogeneity:

 An application to summer jobs. *American Economic Review*, 107(5), 546–550.
- Djebbari, H., & Smith, J. (2008). Heterogeneous impacts in PROGRESA. *Journal of Econometrics*, 145(1), 64–80.
- Farrell, M. H. (2015). Robust inference on average treatment effects with possibly more covariates than observations. *Journal of Econometrics*, 189(1), 1–23.

- Friedman, J., Hastie, T., & Tibshirani, R. (2010). Regularization paths for generalized linear models via coordinate descent. *Journal of Statistical Software*, 33(1), 1–22.
- Gerfin, M., & Lechner, M. (2002). A microeconometric evaluation of the active labour market policy in Switzerland. *Economic Journal*, 112(482), 854–893.
- Grimmer, J., Messing, S., & Westwood, S. J. (2017). Estimating heterogeneous treatment effects and the effects of heterogeneous treatments with ensemble methods. *Political Analysis*, 25(4), 413–434.
- Hahn, P. R., Murray, J. S., & Carvalho, C. (2017). Bayesian regression tree models for causal inference: Regularization, confounding, and heterogeneous effects. (2011),
 1–22. Retrieved from http://arxiv.org/abs/1706.09523
- Hastie, T., Tibshirani, R., & Friedman, J. (2009). The elements of statistical learning Data mining, inference, and prediction (2nd ed.). Springer, New York.
- Hill, J. L. (2011). Bayesian nonparametric modeling for causal inference. *Journal of Computational and Graphical Statistics*, 20(1), 217–240.
- Hirano, K., Imbens, G. W., & Ridder, G. (2003). Efficient estimation of average treatment effects using the estimated propensity score. *Econometrica*, 71(4), 1161–1189.
- Horvitz, D. G., & Thompson, D. J. (1952). A generalization of sampling without replacement from a finite universe. *Journal of the American statistical Association*, 47(260), 663–685.
- Huber, M., Lechner, M., & Mellace, G. (2017). Why do tougher caseworkers increase employment? The role of program assignment as a causal mechanism. *Review of Economics and Statistics*, 99(1), 180–183.
- Huber, M., Lechner, M., & Wunsch, C. (2013). The performance of estimators based on the propensity score. *Journal of Econometrics*, 175(1), 1–21.
- Imai, K., & Ratkovic, M. (2013). Estimating treatment effect heterogeneity in randomized program evaluation. *Annals of Applied Statistics*, 7(1), 443–470.
- Imbens, G. W., & Wooldridge, J. M. (2009). Recent developments in the econometrics of program evaluation. *Journal of Economic Literature*, 47(1), 5–86.
- Johansson, F., Shalit, U., & Sontag, D. (2016). Learning representations for counterfactual inference. In *International conference on machine learning* (pp. 3020–3029).

- Kang, J. D. Y., & Schafer, J. L. (2007). Demystifying double robustness: A comparison of alternative strategies for estimating a population mean from incomplete data. Statistical Science, 22(4), 523–539.
- Knaus, M. C., Lechner, M., & Strittmatter, A. (2017). Heterogeneous employment effects of job search programmes: A machine learning approach. Retrieved from http://arxiv.org/abs/1709.10279
- Künzel, S. R., Sekhon, J. S., Bickel, P. J., & Yu, B. (2017). Meta-learners for estimating heterogeneous treatment effects using machine learning. Retrieved from http://arxiv.org/abs/1706.03461
- Lalive, R., van Ours, J., & Zweimüller, J. (2008). The impact of active labor market programs on the duration of unemployment. *Economic Journal*, 118(525), 235–257.
- Lechner, M. (2018). Penalized causal forests for estimating heterogeneous causal effects.

 Working Paper.
- Lechner, M., & Smith, J. (2007). What is the value added by caseworkers? *Labour Economics*, 14(2), 135–151.
- Lechner, M., & Strittmatter, A. (2017). Practical procedures to deal with common support problems in matching estimation. *Econometric Reviews*, forthcoming.
- Lechner, M., & Wunsch, C. (2013). Sensitivity of matching-based program evaluations to the availability of control variables. *Labour Economics*, 21, 111–121.
- Lee, S., Okui, R., & Whang, Y.-J. (2017). Doubly robust uniform confidence band for the conditional average treatment effect function. *Journal of Applied Econometrics*.
- Leeb, H., & Pötscher, B. M. (2005). Model selection and inference: Facts and fiction.

 Econometric Theory, 21(1), 21–59. doi: 10.1017/S0266466605050036
- Leeb, H., & Pötscher, B. M. (2008). Sparse estimators and the oracle property, or the return of Hodges' estimator. *Journal of Econometrics*, 142(1), 201–211. doi: 10.1016/J.JECONOM.2007.05.017
- McFowland, E., Somanchi, S., & Neill, D. B. (2018). Efficient discovery of heterogeneous treatment effects in randomized experiments via anomalous pattern detection.

 Retrieved from http://arxiv.org/abs/1803.09159
- Murphy, S. (2003). Optimal dynamic treatment regimes. Journal of the Royal Statistical

- Society. Series B (Statistical Methodology), 65(2), 331–366.
- Nie, X., & Wager, S. (2017). Quasi-oracle estimation of heterogeneous treatment effects.

 Retrieved from http://arxiv.org/abs/1712.04912
- Oprescu, M., Syrgkanis, V., & Wu, Z. S. (2018). Orthogonal random forest for heterogeneous treatment effect estimation. Retrieved from http://arxiv.org/abs/1806.03467
- Powers, S., Qian, J., Jung, K., Schuler, A., Shah, N. H., Hastie, T., & Tibshirani, R. (2018). Some methods for heterogeneous treatment effect estimation in high dimensions. Statistics in Medicine, 37(11), 1767–1787.
- Qian, M., & Murphy, S. A. (2011). Performance guarantees for individualized treatment rules. *Annals of statistics*, 39(2), 1180.
- Robins, J. M. (2004). Optimal structural nested models for optimal sequential decisions. , 189–326.
- Robins, J. M., & Rotnitzky, A. (1995). Semiparametric efficiency in multivariate regression Models with missing data. *Journal of the American Statistical Association*, 90(429), 122–129.
- Robinson, P. M. (1988). Root-n-consistent semiparametric regression. Econometrica, 56(4), 931.
- Rolling, C. A., Velez, D., & Yang, Y. (2018). Combining estimates of conditional treatment effects. *Econometric Theory*, forthcoming.
- Rosenbaum, P. R., & Rubin, D. B. (1985). Constructing a control group using multivariate matched sampling methods that incorporate the propensity score. *The American Statistician*, 39(1), 33–38.
- Rubin, D. B. (1974). Estimating causal effects of treatments in randomized and nonrandomized studies. *Journal of Educational Psychology*, 66(5), 688–701.
- Schuler, A., Baiocchi, M., Tibshirani, R., & Shah, N. (2018). A comparison of methods for model selection when estimating individual treatment effects. Retrieved from http://arxiv.org/abs/1804.05146
- Schwab, P., Linhardt, L., & Karlen, W. (2018). Perfect match: A simple method for learning representations for counterfactual inference with neural networks. Retrieved from http://arxiv.org/abs/1810.00656

- Shalit, U., Johansson, F. D., & Sontag, D. (2016). Estimating individual treatment effect:

 Generalization bounds and algorithms. Retrieved from http://arxiv.org/abs/
 1606.03976
- Signorovitch, J. E. (2007). *Identifying informative biological markers in high-dimensional* genomic data and clinical trials (Unpublished doctoral dissertation). PhD thesis, Harvard University.
- Strittmatter, A. (2018). What is the value added by causal machine learning methods in a welfare experiment evaluation? *Working Paper*.
- Su, X., Tsai, C.-L., Wang, H., Nickerson, D. M., & Li, B. (2009). Subgroup analysis via recursive partitioning. *Journal of Machine Learning Research*, 10(Feb), 141–158.
- Taddy, M., Gardner, M., Chen, L., & Draper, D. (2016). A nonparametric Bayesian analysis of heterogeneous treatment effects in digital experimentation. *Journal of Business & Economic Statistics*, 34(4), 661–672.
- Tian, L., Alizadeh, A. A., Gentles, A. J., & Tibshirani, R. (2014). A simple method for estimating interactions between a treatment and a large number of covariates. *Journal of the American Statistical Association*, 109(508), 1517–1532.
- Tibshirani, R. (1996). Regression shrinkage and selection via the Lasso. *Journal of the Royal Statistical Society. Series B (Methodological)*, 58, 267–288.
- Waernbaum, I., & Pazzagli, L. (2017). Model misspecification and bias for inverse probability weighting and doubly robust estimators., 1–25. Retrieved from http://arxiv.org/abs/1711.09388
- Wager, S., & Athey, S. (2018). Estimation and inference of heterogeneous treatment effects using random forests. *Journal of the American Statistical Association*, 113(523), 1228–1242.
- Wendling, T., Jung, K., Callahan, A., Schuler, A., Shah, N. H., & Gallego, B. (2018). Comparing methods for estimation of heterogeneous treatment effects using observational data from health care databases. *Statistics in Medicine*.
- Zhang, B., Tsiatis, A. A., Laber, E. B., & Davidian, M. (2012). A robust method for estimating optimal treatment regimes. *Biometrics*, 68(4), 1010–1018.
- Zhao, Q., Small, D. S., & Ertefaie, A. (2017). Selective inference for effect modification

via the lasso. Retrieved from http://arxiv.org/abs/1705.08020

Appendices

A Data

A.1 Dataset

The data we use includes all individuals who are registered as unemployed at Swiss regional employment agencies in the year 2003. The data contains rich information from different unemployment insurance databases (AVAM/ASAL) and social security records (AHV). This is the standard data used for many Swiss ALMP evaluations (e.g., Gerfin & Lechner, 2002; Lalive, van Ours, & Zweimüller, 2008; Lechner & Smith, 2007). We observe (among others) residence status, qualification, education, language skills, employment history, profession, job position, industry of last job, and desired occupation and industry. The administrative data is linked with regional labour market characteristics, such as the population size of municipalities and the cantonal unemployment rate. The availability of extensive caseworker information and their subjective assessment of the employability of their clients distinguishes our data. Swiss caseworkers employed in the period 2003-2004 were surveyed based on a written questionnaire in December 2004 (see Behncke et al., 2010a, 2010b). The questionnaire contained questions about aims and strategies of the caseworker and the regional employment agency.

In total, 238,902 persons registered as being unemployed in 2003. We only consider the first unemployment registration per individual in 2003. Each registered unemployed person is assigned to a caseworker. In most cases, the same caseworker is responsible for the entire unemployment spell of her/his client. If this is not the case, we focus on the first caseworker to avoid concerns about (rare) endogenous caseworker changes (see Behncke et al., 2010a). We only consider unemployed aged between 24 and 55 years who receive unemployment insurance benefits. We omitted unemployed persons who apply for disability insurance benefits, when the responsible caseworker is not clearly defined, or when their caseworkers did not answer the questionnaire (the response rate is 84%). We drop unemployed foreigners with a residence permit that is valid for less than a year.

Finally, we drop unemployed persons from five regional employment agencies that are not comparable to the other regional employment agencies. This sample is identical to the data used in Huber et al. (2017) and contains 100,120 unemployed persons. We drop further 3,822 observations that participated in alternative treatments. This leaves us with 96,298 to estimate the propensity score. After dropping 342 observations with propensity score below 5% and above 95% and dropping the treated, we are left with 88,844 observations for the simulation. 10,000 are used as validation sample and 78,844 to draw the simulation samples.

A.2 Descriptive statistics

This Appendix provides descriptive statistics of the dataset that is used to build the EMCS. Table A.1 shows the mean and the standard deviation of the outcome, the estimated propensity score, and the variables that are used to estimate the propensity score. Other variables and transformations that are part of the 1,749 covariates for the heterogeneity analysis are omitted. The last column reports standardized differences to illustrate covariate imbalances between treated and controls. Standardized differences normalize the absolute mean difference between two groups by the square root of their mean variance:

Std. Diff. =
$$\frac{|\bar{X}_1 - \bar{X}_0|}{\sqrt{1/2(Var(X_1) + Var(X_0))}} \cdot 100$$
 (25)

We observe that the estimated propensity score is highly imbalanced with a standardized difference of 77, while already a value of 20 is considered as indication for large imbalances (Rosenbaum & Rubin, 1985). This shows that we operate in a setting of high selectivity. The imbalances are mainly driven by differences in the language regions, the previous labor market history and employability of the unemployed.

Figure A.1 shows the distribution of the observed non-treated outcomes that are used in the EMCS. We observe mass points at the bounds of the outcome variable. Nearly 30% of the unemployed find no job at all in the 33 months after the start of the job search program and roughly 10% find a job right away and stay employed for the whole period.

Finally, the absolute correlations of the 105 available covariates are visualized as a

so-called heatmap in Figure A.2. The dark spots indicate that some of the covariates are highly correlated.

Table A.1: Descriptive statistics

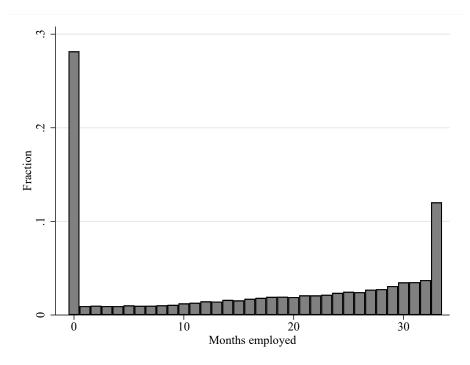
	Trea	ated	Cont	trols	
	Mean	SD	Mean	SD	Std. Diff
Cumulated months of employment: 33 months	16.8	12.8	16.1	12.8	4.8
Propensity score in population	0.11	0.06	0.07	0.05	77.0
Covariates: Characteristics	of unem	ployed p	ersons		
Female	0.46	0.50	0.44	0.50	3.4
*French	0.03	0.16	0.11	0.31	32.1
*Italian	0.01	0.10	0.04	0.18	16.1
Mother tongue other than German, French, Italian	0.26	0.44	0.32	0.47	13.6
*French	0.02	0.14	0.07	0.26	26.1
*Italian	0.01	0.08	0.02	0.15	13.4
Unskilled	0.21	0.40	0.23	0.42	5.5
*French	0.02	0.15	0.05	0.22	13.7
*Italian	0.01	0.10	0.03	0.16	12.4
Qualification: semiskilled	0.13	0.34	0.16	0.37	7.8
*French	0.01	0.11	0.04	0.21	18.9
*Italian	0.002	0.11 0.05	0.04	0.08	6.7
Qualification: skilled without degree	0.002	0.03 0.17	0.01	0.00	8.3
*French	0.002	0.17	0.03 0.02	0.21 0.13	15.8
*Italian	0.002 0.002	0.03 0.04	0.02 0.01	0.13 0.08	6.3
		1.09	$0.51 \\ 0.58$	1.21	7.6
# of unemp. spells in last 2 years	0.49				
*French	0.05	0.40	0.17	0.72	20.0
*Italian	0.02	0.26	0.06	0.42	10.6
Fraction of months emp. in last 2 years	0.84	0.22	0.79	0.25	18.2
*French	0.05	0.20	0.18	0.35	47.7
*Italian	0.02	0.12	0.06	0.21	23.2
Employability rating low	0.11	0.32	0.14	0.34	6.9
*French	0.01	0.08	0.02	0.14	12.4
*Italian	0.004	0.06	0.01	0.10	6.8
Employability rating medium	0.76	0.43	0.74	0.44	3.7
*French	0.05	0.22	0.19	0.39	43.6
*Italian	0.01	0.12	0.04	0.20	17.6
Education: above vocational training	0.45	0.50	0.44	0.50	3.0
*French	0.03	0.17	0.10	0.30	28.6
*Italian	0.01	0.08	0.02	0.15	13.6
Education: tertiary track	0.23	0.42	0.24	0.43	0.6
*French	0.02	0.13	0.09	0.29	31.9
*Italian	0.003	0.06	0.02	0.12	13.0
Vocational training degree	0.28	0.45	0.23	0.42	11.7
*French	0.002	0.05	0.01	0.11	12.1
*Italian	0.01	0.11	0.04	0.19	17.2
Age in 10 year	3.70	0.87	3.66	0.87	5.2
Age squared / 10,000	0.14	0.07	0.14	0.07	5.0
Married	0.45	0.50	0.49	0.50	8.2
Foreigner with B permit	0.10	0.29	0.14	0.35	13.3
Foreigner with C permit	0.21	0.41	0.25	0.43	9.2
Lives in big city	0.14	0.35	0.17	0.37	7.6
Lives in medium sized city	0.14 0.17	0.37	0.13	0.34	9.9
Past income (in CHF 10,000)	0.46	0.20	0.13 0.42	0.34 0.21	20.5
1 act meetin (iii Ciii 10,000)	0.40	0.20			on next pag

Table A.1 – continued from previous page

	Trea	ated	Con	trols	
	Mean	SD	Mean	SD	Std. Diff.
Number of employment spells in last 5 years	0.10	0.13	0.12	0.15	15.1
Previous job in primary sector	0.06	0.24	0.09	0.29	12.3
Previous job in secondary sector	0.15	0.35	0.13	0.34	3.6
Previous job in tertiary sector	0.64	0.48	0.58	0.49	11.5
Foreigner with mother tongue in canton's language	0.12	0.32	0.11	0.32	1.6
Previous job self-employed	0.003	0.06	0.01	0.08	4.5
Previous job manager	0.08	0.28	0.07	0.26	4.1
Previous job skilled worker	0.65	0.48	0.61	0.49	8.4
Previous job unskilled worker	0.24	0.43	0.29	0.45	10.7
Covariates: Allocation of u	inemploye	d to case	workers		
By industry	0.68	0.47	0.54	0.50	30.2
*French	0.03	0.18	0.10	0.29	24.9
*Italian	0.01	0.10	0.03	0.18	16.5
By occupation	0.59	0.49	0.56	0.50	5.2
*French	0.05	0.21	0.17	0.37	39.7
*Italian	0.01	0.11	0.04	0.21	20.3
By age	0.03	0.18	0.03	0.18	0.6
By employability	0.06	0.23	0.07	0.25	4.1
By region	0.09	0.28	0.12	0.32	10.5
Other	0.07	0.25	0.07	0.26	2.9
Covariates: Casewo	rker chara	acteristics			
Age in years	44.68	11.54	44.35	11.60	2.8
*French	2.89	11.60	11.04	20.20	49.5
*Italian	0.95	6.44	3.24	11.61	24.3
Female	0.47	0.50	0.41	0.49	11.5
*French	0.02	0.13	0.09	0.28	32.2
*Italian	0.01	0.10	0.02	0.15	9.8
Tenure (in years)	5.63	3.14	5.83	3.31	6.2
*French	0.37	1.59	1.54	3.03	48.2
*Italian	0.18	1.26	0.57	2.23	21.9
Own unemp. experience	0.63	0.48	0.63	0.48	0.1
*French	0.04	0.20	0.16	0.37	40.7
*Italian	0.02	0.13	0.05	0.21	17.1
Indicator for missing caseworker characteristics	0.04	0.20	0.04	0.20	0.7
Covariates: Local labou	r market o	characteri	stics		
French speaking REA	0.06	0.24	0.24	0.43	52.2
Italian speaking REA	0.02	0.15	0.08	0.26	24.9
Cantonal unemployment rate (in %)	3.55	0.75	3.74	0.86	24.5
*French	0.23	0.93	1.02	1.84	53.9
*Italian	0.10	0.64	0.32	1.13	24.4
Cantonal GDP per capita (in CHF 10,000)	0.51	0.09	0.49	0.09	12.0
Number of observations	7,5	545	88,8	844	

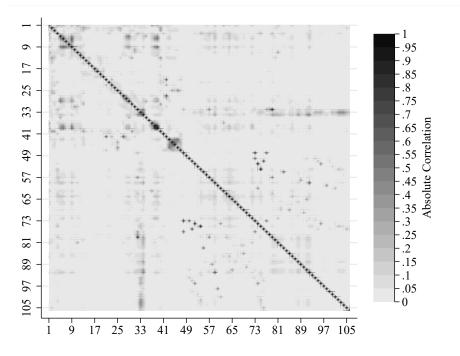
Notes: SD means standard deviation. Std. Diff. stands for standardized differences and are calcualted according to Equation 25.

Figure A.1: Histogram of observed Y^0



Notes: Histogram of the non-treated outcomes counting the months of employment in the 33 months after the start of the job search program.

Figure A.2: Heatmap of baseline covariates



Notes: The heatmap visualizes the absolute correlations of the 105 covariates that are used in the heterogeneity analysis.

B DGP

B.1 Propensity score

Table B.1 shows the estimated propensity score using the specification of Huber et al. (2017). Many variables have sizeable and statistically significant coefficients that create the selection into the treatment. Figure B.1 shows the distribution of the propensity score after the manipulations described in Section 5.2.1. Figure B.1 shows that our setting does not creates problems due to extreme propensity scores or no overlap.

Table B.1: Propensity score

Characteristics of unemployed persons Female 0.140*** 0.00683*** *French -0.0813 -0.00383 *Italian -0.0562 -0.00266 *Italian -0.039** -0.0112*** *French -0.039** -0.0112*** *Trench -0.0312 -0.00150 *French -0.0312 -0.00150 *Italian 0.0151 0.00073 *Italian 0.0151 0.00576** *Italian 0.0151 0.00576** *French (0.0447) (0.0023) *French (0.119*** 0.00576** *Italian (0.144** 0.0259* *Italian (0.441** 0.0259* *Italian (0.436) (0.0140) Qualification: semiskilled (0.13*** -0.00622*** *French (0.838) (0.0436) *Italian (0.63*** 0.0459* *Qualification: skilled without degree (0.0436) (0.0015) *French		(1)	(2)
Female 0.140*** 0.00683*** *French -0.0813 -0.00383 *Italian -0.0562 -0.0266 (0.105) (0.0481) -0.0562 *Italian -0.0562 -0.0266 (0.167) (0.00770) Mother tongue other than German, French, Italian -0.239*** -0.0112*** (0.0393) (0.00177) *French -0.0312 -0.00150 *Italian 0.0151 0.000735 (0.203) (0.00995) Unskilled 0.115*** 0.00576** (0.130) (0.0136) *French 1.019*** 0.0756*** (0.130) (0.0136) *Italian 0.441** 0.0259* *French 0.893*** -0.0622*** *Prench 0.893*** 0.0622*** *Qualification: semiskilled without degree 0.0718 0.0359* *Italian 0.063** 0.00426 *French -0.0561 -0.00265 *French -0.0561 -0.00265		Coefficients	Marginal effects
$\begin{tabular}{l l l l l l l l l l l l l l l l l l l $	Characteristics of unemployed	d persons	
*French	Female	0.140***	0.00683***
		(0.0294)	(0.00145)
*Italian $ \begin{array}{c ccccccccccccccccccccccccccccccccccc$	*French	-0.0813	-0.00383
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$		(0.105)	(0.00481)
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$	*Italian	-0.0562	-0.00266
$\begin{tabular}{l lllllllllllllllllllllllllllllllllll$		(0.167)	(0.00770)
$ \begin{tabular}{l lllllllllllllllllllllllllllllllllll$	Mother tongue other than German, French, Italian	-0.239***	-0.0112***
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$		(0.0393)	(0.00177)
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$	*French	-0.0312	-0.00150
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$		(0.119)	(0.00561)
Unskilled $0.115***$ $0.00576**$ *French (0.0447) (0.00230) *French $1.019***$ $0.0756***$ (0.130) (0.0136) (0.0136) *Italian $0.441**$ $0.0259*$ (0.206) (0.0144) (0.0144) Qualification: semiskilled $-0.134***$ $-0.00622****$ (0.0436) (0.00195) *** *French $0.893***$ $0.0631***$ (0.144) (0.0139) *** *Italian $0.693**$ $0.0459*$ (0.289) (0.0251) Qualification: skilled without degree 0.0718 0.00358 *French -0.0561 -0.00265 *French -0.0561 -0.00265 *Italian 0.313 0.0175 #of unemp. spells in last 2 years 0.0105 0.000511 *French 0.0482 0.00234 *Trench 0.0482 0.00234 *Italian $0.015*$ $0.00510*$	*Italian	` /	` /
Unskilled 0.115^{***} 0.00576^{**} *French 1.019^{***} 0.0756^{***} *Italian 0.441^{**} 0.0259^{*} *Italian 0.441^{**} 0.0259^{*} Qualification: semiskilled -0.134^{***} -0.0062^{***} *French 0.893^{***} 0.0631^{***} *French 0.893^{***} 0.0631^{***} *Italian 0.693^{**} 0.0459^{*} Qualification: skilled without degree 0.0718 0.00358 *French -0.0561 -0.00265 *French -0.0561 -0.00265 *Italian 0.313 0.0175 # of unemp. spells in last 2 years 0.0105 0.000511 *French 0.0482 0.00234 *French 0.0482 0.00234 *Italian $0.015*$ $0.00510*$		(0.203)	(0.00995)
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$	Unskilled		
$ \begin{array}{cccccccccccccccccccccccccccccccccccc$		(0.0447)	(0.00230)
*Italian 0.441^{**} 0.0259^{*} Qualification: semiskilled (0.206) (0.0144) Qualification: semiskilled -0.134^{***} -0.00622^{***} *French 0.893^{***} 0.0631^{***} *Italian 0.693^{**} 0.0459^{*} *Italian 0.693^{**} 0.0459^{*} Qualification: skilled without degree 0.0718 0.00358 *French -0.0561 -0.00265 *French -0.0561 -0.00265 *Italian 0.313 0.0175 # of unemp. spells in last 2 years 0.0105 0.000511 *French 0.0482 0.00234 *French 0.0482 0.00234 *Italian 0.015^{*} 0.00176 *Italian 0.105^{*} 0.00510^{*}	*French		
*Italian 0.441^{**} 0.0259^{*} Qualification: semiskilled (0.206) (0.0144) Qualification: semiskilled -0.134^{***} -0.00622^{***} *French 0.893^{***} 0.0631^{***} *Italian 0.693^{**} 0.0459^{*} *Italian 0.693^{**} 0.0459^{*} Qualification: skilled without degree 0.0718 0.00358 *French -0.0561 -0.00265 *French -0.0561 -0.00265 *Italian 0.313 0.0175 # of unemp. spells in last 2 years 0.0105 0.000511 *French 0.0482 0.00234 *French 0.0482 0.00234 *Italian 0.015^{*} 0.00176 *Italian 0.105^{*} 0.00510^{*}		(0.130)	(0.0136)
$\begin{array}{cccccccccccccccccccccccccccccccccccc$	*Italian		
$\begin{array}{cccccccccccccccccccccccccccccccccccc$		(0.206)	(0.0144)
*French	Qualification: semiskilled		
*French 0.893^{***} 0.0631^{***} (0.144) (0.0139) *Italian 0.693^{**} 0.0459^{*} (0.289) (0.0251) Qualification: skilled without degree 0.0718 0.00358 (0.0828) (0.00426) *French -0.0561 -0.00265 (0.288) (0.0133) *Italian 0.313 0.0175 (0.318) (0.0202) # of unemp. spells in last 2 years 0.0105 0.000511 (0.0131) (0.000634) *French 0.0482 0.00234 (0.0364) (0.00176) *Italian 0.105^{*} 0.00510^{*}	•	(0.0436)	(0.00195)
$ \begin{array}{c} * \text{Italian} & \begin{pmatrix} 0.144 \\ 0.0139 \\ 0.693^{***} \\ 0.0459^{**} \\ 0.289 \\ 0.0251 \\ \end{pmatrix} \\ \text{Qualification: skilled without degree} & \begin{pmatrix} 0.0718 \\ 0.0718 \\ 0.00358 \\ 0.00828 \\ 0.00426 \\ \end{pmatrix} \\ * \text{French} & \begin{pmatrix} -0.0561 \\ -0.00265 \\ 0.288 \\ 0.0133 \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.313 \\ 0.0175 \\ 0.318 \\ 0.0202 \\ \end{pmatrix} \\ \# \text{ of unemp. spells in last 2 years} & \begin{pmatrix} 0.0105 \\ 0.000511 \\ 0.000634 \\ \end{pmatrix} \\ * \text{French} & \begin{pmatrix} 0.0482 \\ 0.00234 \\ 0.00364 \\ 0.00176 \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.105^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**} \\ 0.00510^{**} \\ \end{pmatrix} \\ * \text{Italian} & \begin{pmatrix} 0.00510^{**}$	*French		
*Italian 0.693^{**} 0.0459^{*} (0.289) (0.0251) Qualification: skilled without degree 0.0718 0.00358 (0.0828) (0.00426) *French -0.0561 -0.00265 (0.288) (0.0133) *Italian 0.313 0.0175 (0.318) (0.0202) # of unemp. spells in last 2 years 0.0105 0.000511 *French 0.0482 0.00234 (0.0364) (0.00176) *Italian 0.105^* 0.00510^*		(0.144)	
$\begin{array}{c} \text{Qualification: skilled without degree} & 0.0718 & 0.00358 \\ & (0.0828) & (0.00426) \\ * French & -0.0561 & -0.00265 \\ & (0.288) & (0.0133) \\ * Italian & 0.313 & 0.0175 \\ & (0.318) & (0.0202) \\ \# \ of \ unemp. \ spells \ in \ last \ 2 \ years & 0.0105 & 0.000511 \\ & (0.0131) & (0.000634) \\ * French & 0.0482 & 0.00234 \\ & (0.0364) & (0.00176) \\ * Italian & 0.105* & 0.00510* \\ \end{array}$	*Italian		
$\begin{array}{c} \text{Qualification: skilled without degree} & 0.0718 & 0.00358 \\ & (0.0828) & (0.00426) \\ * \text{French} & -0.0561 & -0.00265 \\ & (0.288) & (0.0133) \\ * \text{Italian} & 0.313 & 0.0175 \\ & (0.318) & (0.0202) \\ \# \ \text{of unemp. spells in last 2 years} & 0.0105 & 0.000511 \\ & (0.0131) & (0.000634) \\ * \text{French} & 0.0482 & 0.00234 \\ & (0.0364) & (0.00176) \\ * \text{Italian} & 0.105* & 0.00510* \\ \end{array}$		(0.289)	(0.0251)
*French	Qualification: skilled without degree	, ,	` /
*French $ \begin{array}{ccccccccccccccccccccccccccccccccccc$	v	(0.0828)	
*Italian	*French	\	,
*Italian $ \begin{array}{ccccccccccccccccccccccccccccccccccc$			
# of unemp. spells in last 2 years	*Italian	\ /	\ /
# of unemp. spells in last 2 years			
*French (0.0131) (0.000634) *French 0.0482 0.00234 (0.0364) (0.00176) *Italian $0.105*$ $0.00510*$	# of unemp. spells in last 2 years		\ /
*French 0.0482 0.00234 (0.0364) (0.00176) *Italian $0.105*$ $0.00510*$,, , , , , , , , , , , , , , , , , , , ,		
Italian (0.0364) (0.00176) 0.105 $0.00510*$	*French	\	,
Italian 0.105 $0.00510*$			
	*Italian		

Table B.1 – continued from previous page

Table B.1 Continued from	1 previous page	(2)
	(1)	(2)
	Coefficients	Marginal effects
	(0.0558)	(0.00270)
Fraction of months emp. in last 2 years	0.225***	0.0109***
	(0.0640)	(0.00310)
*French	-0.108	-0.00522
	(0.199)	(0.00967)
*Italian	-0.0512	-0.00248
	(0.334)	(0.0162)
Employability rating low	-0.641***	-0.0255***
	(0.0562)	(0.00184)
*French	1.323***	0.115***
	(0.248)	(0.0335)
*Italian	1.335***	0.118***
	(0.291)	(0.0404)
Employability rating medium	-0.310***	-0.0161***
	(0.0414)	(0.00232)
*French	0.879***	0.0558***
	(0.194)	(0.0157)
*Italian	0.773***	0.0519***
	(0.224)	(0.0200)
Education: above vocational training	0.0594*	0.00289*
	(0.0311)	(0.00152)
*French	0.115	0.00581
	(0.139)	(0.00733)
*Italian	0.131	0.00674
	(0.198)	(0.0107)
Education: tertiary track	0.167***	0.00843***
v	(0.0380)	(0.00200)
French	-0.285	-0.0125**
	(0.151)	(0.00592)
*Italian	-0.281	-0.0121
	(0.274)	(0.0104)
Vocational training degree	0.120***	0.00601***
Toody training degree	(0.0308)	(0.00158)
*French	-0.332	-0.0139
	(0.267)	(0.00963)
*Italian	0.125	0.00638
	(0.180)	(0.00965)
Age in 10 year	0.0955	0.00463
1180 11 10 9001	(0.136)	(0.00662)
Age squared / 10,000	-0.814	-0.0395
rige squared / 10,000	(1.755)	(0.0851)
Married	-0.0235	-0.00114
Walled	(0.0292)	(0.00114)
Foreigner with B permit	-0.187***	-0.00852***
Totol8hot with D betime	(0.0511)	(0.00312)
Foreigner with C permit	-0.0815**	-0.00388**
roroigner with O berinit	(0.0376)	(0.00176)
Lives in his city	-0.138***	-0.00640***
Lives in big city		
Times in modium sined site	(0.0414)	(0.00185)
Lives in medium sized city	0.223***	0.0117***
Dt : (:- CHE 10 000)	(0.0361)	(0.00203)
Past income (in CHF 10,000)	0.754***	0.0366***
NT 1 (1 , 11 · 1 , 12	(0.0763)	(0.00371)
Number of employment spells in last 5 years	-0.649***	-0.0314***
	Contin	ued on next page

Table B.1 – continued from previous page

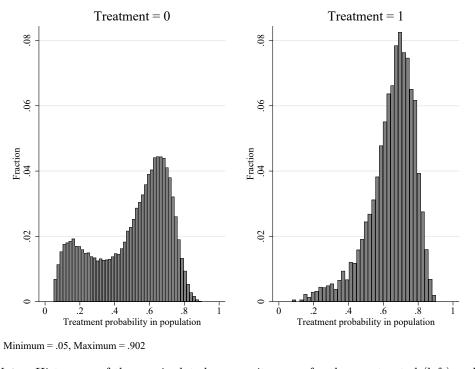
Table B.1 – continued from pre		
	(1)	(2)
	Coefficients	Marginal effects
	(0.105)	(0.00511)
Previous job in primary sector	-0.263***	-0.0116***
	(0.0619)	(0.00247)
Previous job in secondary sector	0.198***	0.0103***
	(0.0481)	(0.00266)
Previous job in tertiary sector	0.113***	0.00545***
	(0.0373)	(0.00178)
Foreigner with mother tongue in canton's language	0.302***	0.0163***
D 1 16 1 1	(0.0430)	(0.00257)
Previous job self-employed	-0.884***	-0.0295***
D : :1	(0.236) -0.480***	(0.00509)
Previous job manager		-0.0195***
Di i.ll.:111l	(0.0970) -0.306***	(0.00326) $-0.0153***$
Previous job skilled worker		
Dravious job unskilled worker	(0.0841) -0.309***	(0.00436) $-0.0141***$
Previous job unskilled worker		
	(0.0896)	(0.00388)
Allocation of unemployed to ca		
By industry	0.349***	0.0167***
	(0.0296)	(0.00142)
*French	0.114	0.00579
	(0.107)	(0.00563)
*Italian	-0.529***	-0.0207***
	(0.175)	(0.00540)
By occupation	0.201***	0.00963***
	(0.0280)	(0.00134)
*French	0.487***	0.0275***
4	(0.122)	(0.00792)
*Italian	-0.463***	-0.0186***
_	(0.171)	(0.00566)
By age	-0.0511	-0.00243
T 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1	(0.0720)	(0.00334)
By employability	-0.362***	-0.0153***
	(0.0549)	(0.00201)
By region	-0.349***	-0.0151***
0.1	(0.0451)	(0.00173)
Other	-0.260***	-0.0114***
	(0.0512)	(0.00204)
Caseworker characterist	cics	
Age (in 10 years)	-0.000912	-4.42e-05
	(0.00136)	(6.60e-05)
*French	0.0217***	0.00105***
	(0.00547)	(0.000264)
*Italian	0.00413	0.000200
	(0.0109)	(0.000528)
Female	0.205***	0.0101***
	(0.0276)	(0.00139)
*French	-0.272**	-0.0119**
	(0.118)	(0.00466)
	والموادوات والمراد	0.0279**
*Italian	0.469***	0.0219
*Italian	(0.172)	(0.0123)
*Italian Tenure (in years)		

Table B.1 – continued from previous page $\,$

	(1)	(2)
	Coefficients	Marginal effects
	(0.00425)	(0.000206)
*French	-0.0749***	-0.00363***
	(0.0196)	(0.000946)
*Italian	0.00549	0.000266
	(0.0276)	(0.00134)
Own unemp. experience	0.0251	0.00121
	(0.0281)	(0.00135)
*French	-0.00642	-0.000311
	(0.115)	(0.00556)
*Italian	0.627***	0.0395**
	(0.197)	(0.0157)
Indicator for missing caseworker characteristics	0.0950	0.00479
Ü	(0.0759)	(0.00398)
Local labour market chara	cteristics	
French speaking REA	-1.346***	-0.0497***
-	(0.496)	(0.0143)
Italian speaking REA	-3.548***	-0.0628***
-	(1.035)	(0.00636)
Cantonal unemployment rate (in %)	0.215***	0.0104***
<u> </u>	(0.0256)	(0.00124)
*French	-0.690***	-0.0334***
	(0.0666)	(0.00314)
*Italian	-0.00787	-0.000381
	(0.179)	(0.00869)
Cantonal GDP per capita (in CHF 10,000)	-3.805***	-0.184***
, , , ,	(0.232)	(0.0113)
Constant	-1.700***	,
	(0.290)	
Number of observations	96,298	96,298

Notes: Coefficients and average marginal effects of the propensity score based on the specification of Huber et al. (2017). Standard errors are in parentheses. ***/**/* indicate statistical significance at the 1%/5%/10%-level.

Figure B.1: Propensity scores by treatment status



Notes: Histogram of the manipulated propensity score for the non-treated (left) and the treated (right).

B.2 Description of ITEs

This Appendix complements the basic statistics of the baseline DGP provided in Table 4 in the main text. Figures B.2 and B.3 show the histograms of the baseline ITE1 and ITE2 with random noise, respectively. We observe a bunching of ITEs at zero because we force the observations to respect the bounds of the outcome variable (see Figure A.1).

Figures B.4 and B.5 compare the cumulative density functions of Y_i^0 and Y_i^1 in the population for the baseline ITE1 and ITE2, respectively. Adding the ITEs to Y_i^0 while respecting the bounds results in more mass of Y_i^1 away from the bounds.³⁰ Further, we observe that the impact of ITE1 on the distribution of the potential outcomes is moderate, while the larger ITE2 changes the potential outcome distribution substantially.

Figure B.6 shows the relation of the propensity score with the ITEs. The 'simulated' caseworkers in our setting are rather successful in identifying unemployed with gains from the program. Those unemployed with a probability of lower than 50% of being send to the program have on average also negative IATEs, which is evident from the dashed lines of the treated potential outcomes being below the solid line of the non-treated potential outcomes. In contrast, the unemployed that participate with a probability of more than 50% benefit on average from the job search program. The simulated assignment mechanism is therefore favorable for most unemployed. However, we build in one feature that is often observed in applications, namely 'cream-skimming' (see, e.g., Bell & Orr, 2002). This means that unemployed persons with good labour market opportunities (a high Y_i^0) have a greater probability to participate in a JSP. However, the effect of the program is not necessarily positive for participants with good labour market opportunities because these participants would have good labour market opportunities even in the absence of training and just suffer from the lock-in effect (see e.g., Card, Kluve, & Weber, 2017). The downward sloping average potential outcomes for very high propensity scores reflect this empirical observations.

Finally, Table B.2 provides an idea about how much of the variation in the potential outcomes, ITEs and the treatment can be explained by the observable characteristics at our disposal. Potential outcomes are relatively hard to approximate. The potential outcome of

 $^{^{30}}$ This is also the reason for the lower standard deviations of the Y_i^1 s in Table 4.

Figure B.2: Histogram of ITE1

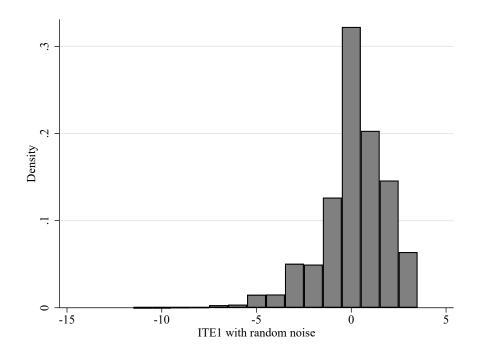
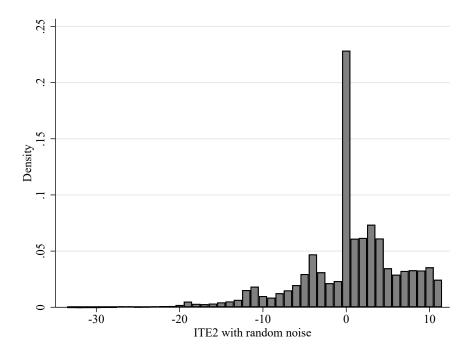


Figure B.3: Histogram of ITE2



non-treated that is taken from real data shows an out-of-sample R^2 of at most about 10%. The ITEs without noise are easier to explain with up to about 70% explainable variation of ITE2. As we add these systematically varying ITE to the non-treated outcomes to obtain the treated outcomes, the different Y_i^1 are easier to approximate compared to Y_i^0 with

explained variation of up to 30%. The picture is changed if we add random noise to create the baseline ITE1 and ITE2. This makes the ITEs and the respective Y_i^1 s much harder to approximate with R^2 of at most 11%. The propensity score is the component that is most easy to approximate. The Lasso is even able to explain 80% of the variation. The reason is that the Lasso has access to all relevant variables and uses the true link function. However, it may not recover exactly the true model due to the shrunken coefficiencts.

Table B.2: Out-of-sample R-squared of different components in %

Sample size:	1000		4000	
Machine learner:	Random Forest	Lasso	Random Forest	Lasso
Without random	noise $(\varepsilon_i = 0)$:			
Y^0	7.6	6.5	10.6	10.7
Y^1 ITE1	9.4	8.2	12.5	13.1
Y^1 ITE2	16.5	22.0	25.2	30.5
ITE1	56.8	62.3	62.1	67.7
ITE2	55.0	59.9	60.2	65.4
Propensity score	56.8	62.9	75.2	80.3
With random nois	se $(\varepsilon_i \sim 1 - Poiss$	son(1)):		
Y^0	7.6	6.5	10.6	10.7
Y^1 ITE0	8.1	6.6	10.9	10.7
Y^1 ITE1	8.4	6.9	11.2	11.1
Y^1 ITE2	7.1	5.8	9.7	10.3
ITE0	0.04	0.02	0.3	0.1
ITE1	4.8	4.2	6.3	6.3
ITE2	3.2	2.7	4.7	4.8
Propensity score	56.8	62.9	75.2	80.3

Notes: Table shows the average out-of-sample R-squared in the validation sample over all replications.

Figure B.4: Cumulative density functions of the potential outcomes of ITE1

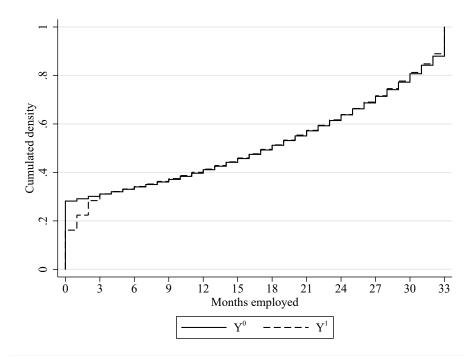


Figure B.5: Cumulative density functions of the potential outcomes of ITE2

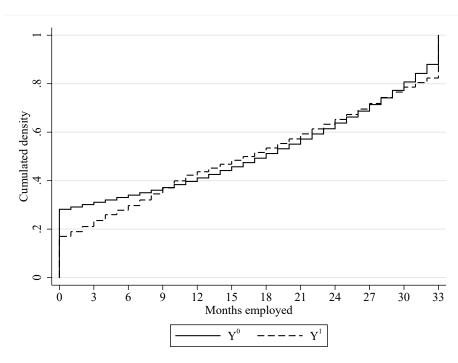
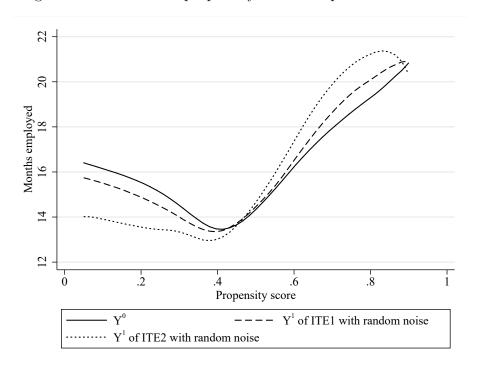


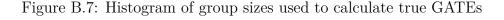
Figure B.6: Relation of propensity score and potential outcomes



Notes: Local constant regression with Epanechnikov kernel and Silverman's bandwidth rule.

B.3 Specification and description of GATEs

We want to create a setting where we need to summarize the heterogeneity over groups that are of interest to the policy maker and might be used by caseworkers to assign program participation. We consider the scenario where we categorize the unemployed by six characteristics: employability, gender, age, qualification, foreigner and language region. This splits the 10,000 validation observations into 64 groups of sizes 32 to 420 (histogram in Figure B.7) by using the following combinations: employability (three categories) x female (binary) x foreigner (binary) x some qualification degree (binary). The biggest group with medium employability is additionally interacted with three age groups (< 30, 30 - 40, > 40) and the German speaking cantons of Switzerland (binary). Figures B.8 and B.9 show the histogram of the distribution of the resulting GATEs of the baseline ITE1 and ITE2.



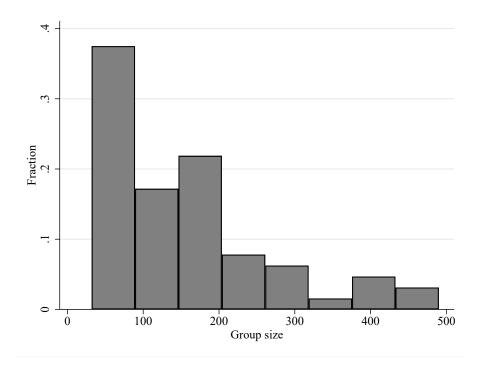


Figure B.8: Histogram of the GATEs of ITE1

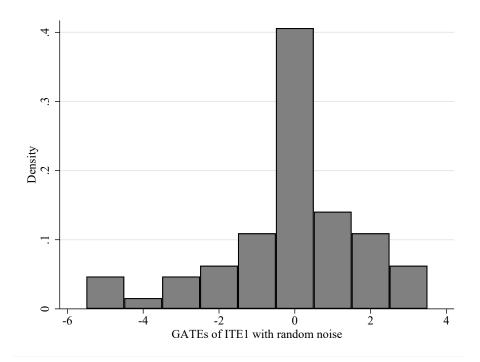
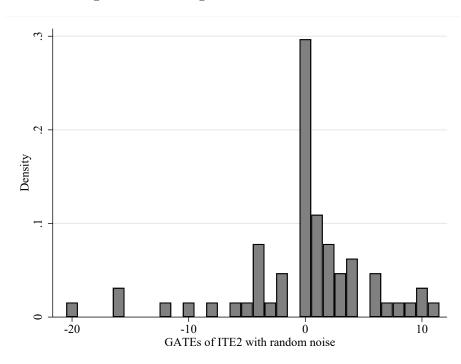


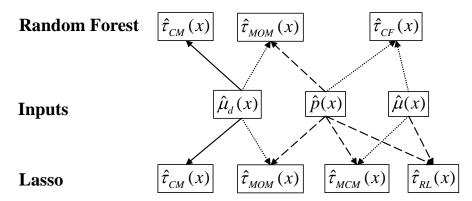
Figure B.9: Histogram of the GATEs of ITE2 $\,$



C Implementation

This Appendix provides a brief description of the implementation steps of the compared estimators.³¹ Before, Figure C.1 provides a graphical summary how different parameters enter the estimation as either the only inputs, as necessary, or as optional nuisance parameters. It shows also how we can use one input for multiple estimators such that the estimation of IATEs requires at most one more step after having $\hat{\mu}_d(x)$, $\hat{p}(x)$ or $\hat{\mu}(x)$. These synergies are important to keep computational time under control.

Figure C.1: Overview of inputs and estimators



Notes: Solid lines indicate that this is the only input. Dashed lines represent necessary nuisance parameters. Dotted lines indicate optional nuisance parameters. $\tau_{MOM}(x)$ summarizes IPW and DR versions of MOM approaches.

Table C.1: Conditional mean regression

- 1. Regress Y_i on X_i in the non-treated sample to obtain a prediction model for $\hat{\mu}(0,x)$.
- 2. Regress Y_i on X_i in the treated sample to obtain a prediction model for $\hat{\mu}(1,x)$.
- 3. Calculate $\hat{\tau}_{CMR}(x) = \hat{\mu}(1,x) \hat{\mu}(0,x)$ for each observation in the validation sample.

Table C.2: Causal Forest

- 1. Apply Algorithm 1 of Athey et al. (2018) with the pseudo outcomes of Equation 17 in the so-called labeling step.
- 2. Calculate $\hat{\tau}_{CF}(x)$ for each observation in the validation sample according to 16 using the weights obtained in step 1.

³¹We do not repeat the tuning parameter choices at each prediction step because they follow always the procedures that are described in Section 4.1 for Random Forest and Section 4.2 for Lasso. Similarly, we omit the cross-fitting steps because the basic principle is already explained in Section 4.3.

Table C.3: Causal Forest with local centering

- 1. Regress D_i on X_i to obtain a prediction model for $\hat{p}(x)$.
- 2. Regress Y_i on X_i to obtain a prediction model for $\hat{\mu}(x)$.
- 3. Apply Algorithm 1 of Athey et al. (2018) with the pseudo outcomes of Equation 17 but replacing D_i and Y_i with $D_i p(X_i)$ and $Y_i \mu(X_i)$.
- 4. Calculate $\hat{\tau}_{CF_LC}(x)$ for each observation in the validation sample according to 16 using the weights obtained in step 3.

Table C.4: MOM with IPW

- 1. Regress D_i on X_i to obtain a prediction model for $\hat{p}(x)$.
- 2. Create the modified outcome $Y_{i,IPW}^*$ by replacing p(x) with $\hat{p}(x)$ in Equation 7.
- 3. Regress $Y_{i,IPW}^*$ on X_i to obtain $\hat{\tau}_{IPW}(x)$ and use it to calculate the IATE for each observation in the validation sample.

Table C.5: MOM with DR

- 1. Regress D_i on X_i to obtain a prediction model for $\hat{p}(x)$.
- 2. Regress Y_i on X_i in the non-treated sample to obtain a prediction model for $\hat{\mu}(0,x)$.
- 3. Regress Y_i on X_i in the treated sample to obtain a prediction model for $\hat{\mu}(1,x)$.
- 4. Create the modified outcome $Y_{i,DR}^*$ by replacing the nuisance parameters by their estimates $\hat{p}(x)$, $\hat{\mu}(0,x)$, and $\hat{\mu}(1,x)$ in Equation 8.
- 5. Regress $Y_{i,DR}^*$ on X_i to obtain $\hat{\tau}_{DR}(x)$ and use it to calculate the IATE for each observation in the validation sample.

The following three tables indicate two different ways to implement the estimators. Either by modifying the covariates (indicated by a) or by modifying the outcomes (indicated by b).

Table C.6: MCM

- 1. Regress D_i on X_i to obtain a prediction model for $\hat{p}(x)$.
- 2a. Modify the covariates as $X_i^{MCM} = T_i/2X_i$.
- 2b. Modify the outcome as $Y_{MCM}^* = 2T_iY_i$.
- 3a. Use $\hat{p}(x)$ and X_i^{MCM} to obtain $\hat{\beta}_{MCM}$ from Equation 10.
- 3b. Use $\hat{p}(x)$ and Y_{MCM}^* to obtain $\hat{\beta}_{MCM}$ from Equation 11 with $\tau(x) = x\beta$.
- 4. Calculate the IATE for each observation in the validation sample as $\hat{\tau}_{MCM}(x) = x\hat{\beta}_{MCM}$.

Table C.7: MCM with efficiency augmentation

- 1. Regress D_i on X_i to obtain a prediction model for $\hat{p}(x)$.
- 2. Regress Y_i on X_i to obtain a prediction model for $\hat{\mu}(x)$.
- 3a. Modify the covariates as $X_i^{MCM} = T_i/2X_i$.
- 3b. Modify the outcome as $Y_{MCMEA}^* = 2T_i(Y_i \mu(X_i))$.
- 4a. Use $\hat{p}(x)$ and X_i^{MCM} to obtain $\hat{\beta}_{MCM_EA}$ from Equation 10 with Y_i being replaced by $Y_i \hat{\mu}(x)$.
- 4b. Use $\hat{p}(x)$ and $Y_{MCM_EA}^*$ to obtain $\hat{\beta}_{MCM_EA}$ from Equation 11 with $\tau(x) = x\beta$.
 - 5. Calculate the IATE for each observation in the validation sample as $\hat{\tau}_{MCM_EA}(x) = x \hat{\beta}_{MCM_EA}$.

Table C.8: R-learning

- 1. Regress D_i on X_i to obtain a prediction model for $\hat{p}(x)$.
- 2. Regress Y_i on X_i to obtain a prediction model for $\hat{\mu}(x)$.
- 3a. Modify the covariates as $X_i^{RL} = (D_i p(X_i))X$.
- 3b. Modify the outcome as $Y_{RL}^* = \frac{Y_i \mu(X_i)}{D_i p(X_i)}$
- 4a. Use $\hat{p}(x)$, $\hat{\mu}(x)$ and X_i^{RL} to obtain $\hat{\beta}_{RL}$ from Equation 13.
- 4b. Use $\hat{p}(x)$, $\hat{\mu}(x)$ and Y_{RL}^* to obtain $\hat{\beta}_{RL}$ from Equation 14 with $\tau(x) = x\beta$.
 - 5. Calculate the IATE for each observation in the validation sample as $\hat{\tau}_{RL}(x) = x\hat{\beta}_{RL}$.

D More results

This Appendix provides the results for all settings with the full set of performance measures. Additionally to the basic measure of Section 5.2.3, we calculate and report the following performance measures for IATEs and GATEs.

To understand simulation noise, we calculate the standard error of our main measure, \overline{MSE} , by

$$SE(\overline{MSE}) = \sqrt{\frac{1}{R} \sum_{r=1}^{R} [MSE_r - \overline{MSE}]^2},$$
(26)

where $MSE_r = \frac{1}{N_v} \sum_{v=1}^{N_v} [\xi(x_v, y_v^0) - \hat{\tau}(x_v)_r]^2$. This measure indicates how precise the mean MSE is measured and is used to assess whether the performance differs significantly or

³²This can also be regarded as the standard error of the mean Precision in Estimation of Heterogenous Effect (PEHE) introduced by Hill (2011). Note that the mean MSE and mean PEHE result in the same number.

just up to noise. Furthermore, instead of taking the mean over all validation observations v, we consider the median $(Median(MSE_v))$. This measure leads sometimes to different orderings of the estimators compared to \overline{MSE} , which indicates that outliers play a role in measuring the performance.

Additionally to the mean absolute bias, we report the mean bias:

$$\overline{Bias} = \frac{1}{N_v} \sum_{v=1}^{N_v} \left[\frac{1}{R} \sum_{r=1}^{R} \hat{\tau}(x_v)_r - \xi(x_v, y_v^0) \right]$$
 (27)

Additionally to the fraction of rejected JB tests, we report the mean skewness and mean kurtosis over the validation sample.

All previous measures summarize the performance over all individuals. However, the measures in the last two columns in the following tables summarize the performance on the replication level. One is the correlation between the estimated and the true ITEs (Corr.) and one is the variance ratio of the estimated and the true ITEs (Var. Ratio),

$$Corr. = \frac{1}{R} \sum_{r=1}^{R} \rho_{\hat{\tau}_r, \xi}, \tag{28}$$

where $\rho_{X,Y}$ denotes the correlation between two variables and $\hat{\tau}_r$ and ξ are vectors of length N_v containing the estimated ITEs in replication r and the true ITEs, respectively.

Var. Ratio =
$$\frac{1}{R} \sum_{r=1}^{R} \frac{Var(\hat{\tau}_r)}{Var(\xi)},$$
 (29)

where $Var(\cdot)$ is the variance of the respective vector.

D.1 Results for IATE estimation

The Appendices D.1.1 to D.1.4 show the full results for IATE estimation in the 24 DGP-sample size combinations. The correlations as additional performance measure in columns 10 show that the estimated IATEs are mostly positively correlated with the true ITEs. This shows that the considered estimators find systematic variation. However, the size of the correlations vary for the different settings. They are larger for the DGP without noise and for larger sample sizes, as expected.

The variance ratios in columns 11 show that estimators tend to overshoot and to create IATEs that vary more than the true ITEs. This is reflected in variance ratios of above one that are particularly prevalent for ITE0 and ITE1. This suggests that estimators tend to overshoot in settings with little variation of the ITEs relative to the variation in the outcome. Especially, both MOM IPW estimators and MCM are prone to create highly variable IATEs. This is in line with the high mean SD of these methods that use only the inverse propensity score to correct for selection bias.

D.1.1 ITE with selection and without random noise

Table D.1 provides additional information for the baseline ITE0. The alternative performance measures confirm the results in the main text. Furthermore, note that the substantial positive mean bias suggests that the estimators are not able to completely remove the positive selection bias which is created by the positive relation between propensity score and ITEs.

The results for ITE1 without random noise are very similar to the baseline results for ITE1 (Table D.2). However, ITE2 without random noise shows some notable differences to its baseline version with random noise. Random Forest based MOM IPW seems to work quite well for ITE2 as it shows the lowest mean MSE (Table D.3). However, the median MSE in column three suggests that this is driven by some outliers because the outstanding performance is not observed for this measure.

Table D.1: Performance measure for ITE0 with selection and without random noise (baseline)

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	ervatio	ns				
Random Forest:											
Infeasible			N	o variati	on in de	pender		ole			
Conditional mean regression	3.69	0.04	3.60	0.62	0.60	1.78	6%	0.0	3.0	-	-
MOM IPW	10.52	0.08	8.52	2.05	0.71	2.16	18%	0.0	3.1	-	-
MOM DR	2.00	0.02	1.94	0.40	0.40	1.35	7%	0.0	3.0	-	-
Causal Forest	3.52	0.04	3.44	0.75	0.75	1.69	12%	0.0	3.1	-	-
Causal Forest with local centering	3.42	0.03	3.29	0.34	0.34	1.81	10%	0.0	3.0	-	-
Lasso:											
Infeasible			N	o variati	on in de	pender	nt variab	ole			
Conditional mean regression	11.21	0.11	10.28	0.69	0.60	3.19	91%	-0.1	3.9	-	-
MOM IPW	11.31	0.28	9.69	1.09	0.61	2.99	100%	0.7	18.7	-	-
MOM DR	45.39	43.11	37.31	0.60	0.60	6.31	100%	41.1	1794.2	-	-
MCM	13.03	0.27	10.46	1.50	0.45	3.05	100%	-0.1	7.5	-	_
MCM with efficiency augmentation	2.08	0.05	1.91	0.42	0.42	1.36	100%	0.0	8.7	-	_
R-learning	2.03	0.05	1.85	0.45	0.45	1.33	100%	0.0	9.4	-	-
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible			N	o variati	on in de	pender	nt variab	ole			
Conditional mean regression	2.79	0.04	2.64	0.61	0.53	1.49	4%	0.0	3.0	-	-
MOM IPW	5.76	0.06	4.44	1.31	0.53	1.74	12%	0.0	3.1	-	-
MOM DR	1.16	0.01	1.10	0.28	0.28	1.03	8%	0.0	3.0	-	-
Causal Forest	2.31	0.03	2.24	0.72	0.70	1.29	6%	0.0	3.0	-	_
Causal Forest with local centering	2.05	0.02	1.94	0.24	0.24	1.40	9%	0.0	3.0	-	-
Lasso:											
Infeasible			N	o variati	on in de	pender	nt variab	ole			
Conditional mean regression	6.14	0.08	5.62	0.65	0.49	2.30	35%	0.0	3.4	-	_
MOM IPW	5.02	0.17	4.31	0.82	0.45	1.93	100%	0.2	7.9	_	-
MOM DR	0.51	0.02	0.43	0.31	0.31	0.62	35%	0.0	3.4	_	-
MCM	5.79	0.22	4.28	1.03	0.34	1.94	100%	0.0	6.5	_	-
MCM with efficiency augmentation	0.48	0.02	0.42	0.26	0.26	0.62	97%	0.0	8.7	_	-
R-learning	0.47	0.02	0.42	0.28	0.28	0.61	97%	0.1	8.7	_	-

Table D.2: Performance measure for ITE1 with selection and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	
		1000 observations										
Random Forest:												
Infeasible	1.34	0.00	0.92	0.98	0.00	0.14	59%	0.0	3.1	0.76	0.48	
Conditional mean regression	7.22	0.05	4.86	1.54	1.16	1.81	7%	0.0	3.0	0.14	0.98	
MOM IPW	10.20	0.08	7.35	1.80	1.16	2.24	17%	0.0	3.1	0.60	4.03	
MOM DR	4.80	0.03	2.88	1.37	0.75	1.33	11%	0.0	3.0	0.18	0.45	
Causal Forest	7.21	0.05	4.65	1.63	1.36	1.70	12%	0.0	3.1	0.16	0.82	
Causal Forest with local centering	6.22	0.04	4.43	1.35	0.72	1.79	12%	0.0	3.0	0.13	0.88	
Lasso:												
Infeasible	1.17	0.00	0.68	0.83	0.01	0.33	67%	0.0	3.5	0.79	0.60	
Conditional mean regression	14.64	0.12	13.09	1.41	1.01	3.30	90%	-0.1	4.3	0.15	3.76	
MOM IPW	13.57	0.30	11.72	1.21	0.94	3.24	100%	0.3	18.0	0.39	4.30	
MOM DR	44.64	39.07	34.97	1.48	0.87	6.03	100%	40.5	1760.1	0.10	2.04	
MCM	13.08	0.22	11.18	1.11	0.57	3.19	100%	-0.1	6.6	0.42	4.54	
MCM with efficiency augmentation	5.31	0.06	3.04	1.42	0.70	1.41	100%	-0.1	9.4	0.11	0.46	
R-learning	5.35	0.06	2.83	1.47	0.79	1.34	100%	0.0	10.3	0.08	0.39	
				400	00 obse	ervatio	ns					
Random Forest:												
Infeasible	1.17	0.00	0.73	0.90	0.00	0.12	22%	0.0	3.0	0.79	0.54	
Conditional mean regression	5.85	0.05	3.81	1.47	1.04	1.51	4%	0.0	3.0	0.21	0.88	
MOM IPW	6.16	0.06	4.48	1.30	0.82	1.80	10%	0.0	3.0	0.55	2.45	
MOM DR	3.50	0.02	2.10	1.24	0.50	1.02	11%	0.0	3.1	0.29	0.36	
Causal Forest	5.68	0.05	3.30	1.55	1.29	1.31	7%	0.0	3.0	0.25	0.64	
Causal Forest with local centering	4.50	0.03	3.07	1.25	0.55	1.40	10%	0.0	3.0	0.21	0.65	
Lasso:												
Infeasible	1.00	0.00	0.50	0.77	0.01	0.23	17%	0.0	3.1	0.82	0.66	
Conditional mean regression	8.41	0.08	7.55	1.20	0.73	2.41	33%	-0.1	3.3	0.31	2.46	
MOM IPW	6.94	0.16	5.92	1.01	0.67	2.21	98%	0.1	6.1	0.49	2.51	
MOM DR	3.45	0.03	1.95	1.36	0.47	0.72	99%	-0.3	9.3	0.24	0.14	
MCM	6.42	0.15	5.00	0.96	0.39	2.09	100%	-0.1	5.6	0.46	2.33	
MCM with efficiency augmentation	3.40	0.02	2.05	1.35	0.43	0.72	99%	-0.1	7.8	0.24	0.14	
R-learning	3.56	0.03	1.87	1.39	0.53	0.70	98%	-0.1	8.2	0.19	0.12	

Table D.3: Performance measure for ITE2 with selection and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Infeasible	18.84	0.01	12.79	3.66	-0.04	0.54	39%	0.0	3.1	0.74	0.47
Conditional mean regression	37.71	0.11	17.19	4.55	2.24	1.89	20%	-0.1	3.0	0.50	0.15
MOM IPW	30.30	0.09	19.54	4.10	2.25	2.37	17%	0.0	3.1	0.66	0.59
MOM DR	37.78	0.07	12.38	4.61	1.64	1.36	29%	0.0	3.1	0.47	0.06
Causal Forest	43.79	0.12	16.98	4.94	2.85	1.77	35%	-0.1	3.1	0.41	0.10
Causal Forest with local centering	37.05	0.09	14.05	4.46	1.59	1.90	36%	0.0	3.1	0.45	0.12
Lasso:											
Infeasible	16.81	0.01	10.78	3.22	-0.03	1.23	71%	0.0	3.4	0.77	0.57
Conditional mean regression	37.10	0.12	26.65	3.93	1.46	3.68	84%	-0.1	3.8	0.52	0.64
MOM IPW	38.53	0.29	28.54	3.86	1.77	3.95	100%	-0.1	10.4	0.55	0.71
MOM DR	78.97	39.21	47.70	4.54	1.32	6.25	100%	37.1	1574.6	0.42	0.27
MCM	36.49	0.15	25.60	3.91	0.97	3.60	100%	-0.1	5.7	0.50	0.53
MCM with efficiency augmentation	38.48	0.09	14.41	4.55	1.30	1.91	100%	-0.1	6.7	0.41	0.11
R-learning	40.18	0.10	14.21	4.68	1.53	1.80	100%	0.0	7.7	0.38	0.09
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible	16.68	0.01	11.34	3.43	-0.04	0.45	15%	0.0	3.0	0.78	0.52
Conditional mean regression	27.75	0.09	16.09	4.08	1.62	1.58	6%	0.0	3.0	0.64	0.31
MOM IPW	24.08	0.07	14.71	3.69	1.47	1.92	10%	0.0	3.0	0.70	0.50
MOM DR	27.19	0.07	11.84	4.14	0.97	1.15	12%	0.0	3.0	0.66	0.18
Causal Forest	34.53	0.15	16.01	4.46	2.47	1.51	17%	-0.1	3.0	0.60	0.17
Causal Forest with local centering	25.17	0.08	14.39	3.96	0.97	1.61	10%	0.0	3.0	0.67	0.29
Lasso:											
Infeasible	14.49	0.01	8.35	3.00	-0.03	0.87	16%	0.0	3.1	0.81	0.63
Conditional mean regression	24.47	0.07	17.85	3.48	0.84	2.58	19%	0.0	3.1	0.68	0.69
MOM IPW	25.33	0.12	18.22	3.47	1.18	2.71	76%	-0.1	4.2	0.68	0.68
MOM DR	25.62	0.08	12.79	4.02	0.62	1.30	85%	-0.1	4.9	0.66	0.23
MCM	27.86	0.12	17.80	3.78	0.62	2.50	95%	-0.1	4.4	0.61	0.38
MCM with efficiency augmentation	26.79	0.08	12.27	4.09	0.71	1.30	88%	0.0	4.4	0.65	0.20
R-learning	27.77	0.09	12.04	4.13	0.87	1.32	92%	0.0	4.9	0.64	0.18

D.1.2 ITE with selection and random noise

ITE0 with random noise (Table D.4) shows the same pattern as the baseline ITE0 without noise, only that the levels of mean MSE and mean bias are higher. This is expected because this ITE0 consists mainly of irreducible noise.

Tables D.5 and D.6 provide additional information for the baseline ITE1 and ITE2, respectively. The alternative performance measures confirm the results in the main text of ITE1. However, Random Forest conditional mean regression and Causal Forest show highly competitive median MSE for both sample sizes, which is in contrast to their mean MSE.

Table D.4: Performance measure for ITE0 with selection and random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
	1000 observations										
Random Forest:											
Infeasible	0.79	0.00	0.82	0.62	0.00	0.05	83%	-0.2	3.2	0.04	0.00
Conditional mean regression	4.50	0.04	3.75	0.85	0.60	1.78	6%	0.0	3.0	-0.01	3.47
MOM IPW	11.62	0.08	8.84	2.17	0.71	2.18	17%	0.0	3.1	-0.03	12.11
MOM DR	2.74	0.02	2.11	0.75	0.39	1.33	8%	0.0	3.0	0.00	1.65
Causal Forest	4.30	0.04	3.47	0.90	0.74	1.69	13%	0.0	3.1	-0.01	2.96
Causal Forest with local centering	4.16	0.03	3.59	0.73	0.33	1.80	11%	0.0	3.0	0.00	3.34
Lasso:											
Infeasible	0.79	0.00	0.86	0.62	0.00	0.05	100%	-1.3	21.3	0.02	0.00
Conditional mean regression	11.85	0.11	10.88	0.90	0.60	3.17	90%	-0.1	4.0	-0.01	12.78
MOM IPW	12.20	0.30	10.45	1.25	0.60	2.99	100%	0.7	19.8	0.00	13.12
MOM DR	1346.38	1343.32	24.83	0.87	0.63	7.92	100%	39.5	1703.1	0.00	1636.00
MCM	13.79	0.26	11.26	1.57	0.42	3.05	100%	-0.1	7.4	0.00	15.41
MCM with efficiency augmentation	2.83	0.06	2.17	0.76	0.42	1.34	100%	0.0	9.0	0.00	1.50
R-learning	2.80	0.05	2.11	0.77	0.45	1.32	100%	0.0	9.7	0.00	1.45
				400	0 obsei	rvatio	ıs				
Random Forest:											
Infeasible	0.78	0.00	0.78	0.62	0.00	0.04	30%	-0.2	3.1	0.06	0.00
Conditional mean regression	3.58	0.03	2.86	0.86	0.53	1.49	4%	0.0	3.0	-0.01	2.97
MOM IPW	6.71	0.06	4.89	1.50	0.53	1.76	12%	0.0	3.1	-0.02	6.83
MOM DR	1.92	0.01	1.39	0.71	0.27	1.02	7%	0.0	3.0	0.00	1.19
Causal Forest	3.09	0.03	2.26	0.89	0.70	1.29	6%	0.0	3.0	-0.01	2.10
Causal Forest with local centering	2.81	0.02	2.31	0.70	0.23	1.39	9%	0.0	3.0	0.00	2.32
Lasso:											
Infeasible	0.78	0.00	0.82	0.62	0.00	0.04	100%	-0.1	7.1	0.04	0.00
Conditional mean regression	6.84	0.08	6.12	0.88	0.49	2.28	35%	0.0	3.4	-0.01	7.17
MOM IPW	5.81	0.17	4.81	1.03	0.45	1.92	100%	0.2	7.8	0.00	5.91
MOM DR	1.31	0.02	0.69	0.72	0.31	0.63	94%	-0.1	11.1	0.00	0.33
MCM	6.34	0.20	4.64	1.16	0.31	1.90	100%	-0.1	6.4	0.00	6.75
MCM with efficiency augmentation	1.26	0.02	0.73	0.71	0.26	0.62	96%	0.0	8.3	0.00	0.31
R-learning	1.26	0.02	0.72	0.71	0.28	0.62	96%	0.1	8.1	0.00	0.30

Table D.5: Performance measure for ITE1 with selection and random noise (baseline)

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				10	00 obse	ervatio	ns				
Random Forest:											
Infeasible	2.98	0.00	1.03	1.29	0.01	0.15	71%	-0.2	3.2	0.23	0.03
Conditional mean regression	7.04	0.04	4.58	1.45	0.80	1.78	8%	0.0	3.0	0.02	0.88
MOM IPW	12.92	0.08	8.65	2.26	0.87	2.20	16%	0.0	3.1	0.18	3.34
MOM DR	5.08	0.02	2.92	1.36	0.52	1.33	8%	0.0	3.0	0.03	0.42
Causal Forest	6.86	0.04	4.23	1.49	0.96	1.68	12%	0.0	3.1	0.03	0.75
Causal Forest with local centering	6.50	0.03	4.48	1.35	0.48	1.79	12%	0.0	3.1	0.02	0.83
Lasso:											
Infeasible	3.00	0.00	1.08	1.28	0.01	0.21	100%	-0.5	7.7	0.21	0.04
Conditional mean regression	14.26	0.11	11.96	1.46	0.76	3.16	90%	-0.1	4.1	0.02	3.23
MOM IPW	15.69	1.46	11.75	1.56	0.73	3.12	100%	0.1	30.3	0.12	3.99
MOM DR	48.76	43.14	38.41	1.40	0.70	6.32	100%	40.9	1783.4	0.01	2.17
MCM	15.31	0.26	12.32	1.72	0.46	3.10	100%	-0.1	7.1	0.14	4.10
MCM with efficiency augmentation	5.27	0.06	3.14	1.37	0.52	1.36	100%	-0.1	9.2	0.02	0.40
R-learning	5.16	0.05	2.89	1.38	0.58	1.29	100%	0.0	9.5	0.01	0.34
				400	00 obse	ervatio	ns				
Random Forest:											
Infeasible	2.93	0.00	1.27	1.30	0.01	0.11	26%	-0.1	3.1	0.25	0.06
Conditional mean regression	6.05	0.04	3.63	1.44	0.73	1.49	4%	0.0	3.0	0.02	0.76
MOM IPW	8.42	0.06	5.38	1.76	0.63	1.77	11%	0.0	3.0	0.15	1.92
MOM DR	4.17	0.02	2.02	1.32	0.35	1.01	9%	0.0	3.0	0.05	0.30
Causal Forest	5.61	0.04	3.01	1.48	0.91	1.29	6%	0.0	3.0	0.05	0.54
Causal Forest with local centering	5.10	0.02	3.08	1.32	0.35	1.39	10%	0.0	3.0	0.03	0.59
Lasso:											
Infeasible	2.93	0.00	1.08	1.28	0.01	0.16	83%	-0.1	4.0	0.25	0.06
Conditional mean regression	9.20	0.08	7.16	1.43	0.60	2.30	36%	0.0	3.4	0.04	1.87
MOM IPW	8.03	0.17	5.81	1.46	0.53	2.01	100%	0.1	7.4	0.13	1.75
MOM DR	3.66	0.02	1.31	1.34	0.37	0.64	96%	-0.3	10.1	0.03	0.09
MCM	8.14	0.19	5.44	1.51	0.32	1.96	100%	-0.1	6.2	0.15	1.90
MCM with efficiency augmentation	3.62	0.02	1.22	1.33	0.33	0.63	98%	0.0	8.7	0.03	0.08
R-learning	3.65	0.02	1.07	1.34	0.38	0.63	98%	0.1	8.6	0.02	0.08

Table D.6: Performance measure for ITE2 with selection and random noise (baseline)

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Infeasible	38.46	0.01	8.96	4.43	0.09	0.52	66%	-0.2	3.2	0.19	0.02
Conditional mean regression	43.74	0.05	12.15	4.58	1.24	1.74	8%	0.0	3.0	0.04	0.07
MOM IPW	46.83	0.08	19.41	4.83	1.23	2.23	16%	0.0	3.1	0.17	0.31
MOM DR	41.45	0.03	12.25	4.50	0.82	1.32	10%	0.0	3.0	0.05	0.03
Causal Forest	43.87	0.06	11.76	4.61	1.43	1.66	12%	0.0	3.1	0.04	0.06
Causal Forest with local centering	42.84	0.04	13.72	4.50	0.81	1.78	12%	0.0	3.0	0.04	0.07
Lasso:											
Infeasible	38.66	0.01	8.89	4.42	0.08	0.71	100%	-0.6	8.3	0.18	0.03
Conditional mean regression	50.11	0.11	22.97	4.52	1.12	3.15	92%	-0.1	4.1	0.04	0.26
MOM IPW	49.82	0.25	24.14	4.50	1.02	3.20	100%	0.5	13.9	0.11	0.33
MOM DR	537.16	494.76	21.20	4.55	0.92	5.04	100%	31.9	1310.9	0.03	11.94
MCM	49.25	0.22	23.98	4.47	0.52	3.18	100%	-0.1	6.8	0.13	0.35
MCM with efficiency augmentation	41.99	0.06	12.63	4.51	0.79	1.41	100%	0.0	9.2	0.04	0.04
R-learning	42.13	0.06	11.73	4.54	0.89	1.35	100%	0.1	10.1	0.03	0.03
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible	37.86	0.00	6.30	4.43	0.07	0.41	31%	-0.2	3.1	0.22	0.04
Conditional mean regression	42.26	0.05	11.94	4.54	1.12	1.46	5%	0.0	3.0	0.06	0.06
MOM IPW	42.69	0.06	13.91	4.54	0.88	1.80	11%	0.0	3.0	0.16	0.18
MOM DR	40.03	0.02	12.21	4.45	0.56	1.03	11%	0.0	3.1	0.10	0.03
Causal Forest	42.34	0.05	11.14	4.58	1.36	1.29	7%	0.0	3.0	0.07	0.05
Causal Forest with local centering	41.05	0.03	12.80	4.46	0.63	1.40	9%	0.0	3.0	0.07	0.05
Lasso:											
Infeasible	37.84	0.00	6.68	4.40	0.07	0.53	84%	-0.1	4.2	0.22	0.04
Conditional mean regression	44.31	0.07	16.52	4.46	0.82	2.33	34%	-0.1	3.3	0.10	0.17
MOM IPW	43.21	0.15	15.02	4.43	0.74	2.17	97%	0.1	5.8	0.14	0.19
MOM DR	40.11	0.03	12.55	4.48	0.54	0.76	97%	-0.2	9.9	0.08	0.01
MCM	42.63	0.16	14.91	4.41	0.34	2.07	100%	-0.1	5.6	0.14	0.18
MCM with efficiency augmentation	40.04	0.03	12.45	4.47	0.51	0.75	99%	0.0	7.5	0.08	0.01

D.1.3 ITE with random assignment and without random noise

Table D.7 provides the results for the baseline ITE0 with random assignment. The relative performance order remains unchanged. However, the striking difference is that the mean absolute bias and the mean bias are both close to zero for all estimators. This shows that the interpretation of the results for DGPs with selectivity in D.1.1 is correct and that the observed biases are driven by remaining selection bias.

The comparison of ITE1 with its selective equivalents shows that the remaining mean absolute bias is lower because of the absence of selection bias, which is also reflected in the close to zero mean bias. However, the mean absolute bias remains substantial, which suggest that approximation errors and the irreducible noise play a significant role. The relative performance of the estimators is similar for ITE1. This is not true for ITE2 where conditional mean regression based on Random Forests show the lowest mean MSE. This suggests that settings with large and informative ITEs favor conditional mean regression because the systematic part in the outcomes can be exploited. However, this is the least realistic setting and this locally good performance of conditional mean regression is not confirmed for ITE2 with random noise in the next section.

Table D.7: Performance measure for ITE0 with random assignment and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Infeasible			No	o variatio	on in de	penden	t variab	le			
Conditional mean regression	3.25	0.03	3.15	0.03	0.01	1.80	6%	0.0	3.0	-	-
MOM IPW	4.04	0.03	3.84	0.03	0.01	1.99	10%	0.0	3.1	-	-
MOM DR	1.88	0.02	1.81	0.02	0.01	1.37	6%	0.0	3.0	-	-
Causal Forest	2.92	0.03	2.80	0.02	0.01	1.70	10%	0.0	3.1	-	-
Causal Forest with local centering	2.97	0.02	2.83	0.03	0.02	1.71	11%	0.0	3.1	_	-
Lasso:											
Infeasible			No	o variatio	on in de	penden	t variab	le			
Conditional mean regression	10.40	0.11	9.48	0.06	0.01	3.16	87%	0.0	4.0	-	-
MOM IPW	3.83	0.14	3.29	0.03	0.01	1.90	100%	-0.1	19.4	-	-
MOM DR	1.65	0.05	1.52	0.02	0.01	1.27	100%	0.0	9.4	_	-
MCM	3.62	0.11	3.09	0.03	0.01	1.84	100%	0.1	15.4	_	_
MCM with efficiency augmentation	1.58	0.04	1.43	0.02	0.02	1.24	100%	0.0	8.4	_	_
R-learning	1.55	0.04	1.42	0.02	0.02	1.23	100%	0.0	8.1	-	-
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible			No	o variatio	on in de	penden	t variab	le			
Conditional mean regression	2.29	0.03	2.20	0.06	0.02	1.50	6%	0.0	3.0	-	-
MOM IPW	2.47	0.02	2.28	0.06	0.03	1.55	7%	0.0	3.0	-	-
MOM DR	0.99	0.01	0.94	0.04	0.02	0.99	6%	0.0	3.0	-	_
Causal Forest	1.68	0.01	1.59	0.05	0.03	1.29	7%	0.0	3.0	-	-
Causal Forest with local centering	1.73	0.01	1.62	0.05	0.02	1.31	8%	0.0	3.0	-	_
Lasso:											
Infeasible			No	o variatio	on in de	penden	t variab	le			
Conditional mean regression	5.64	0.07	5.16	0.09	0.02	2.31	32%	0.0	3.3	_	-
MOM IPW	0.80	0.05	0.67	0.04	0.03	0.87	100%	0.0	12.6	_	-
MOM DR	0.35	0.02	0.32	0.03	0.02	0.59	32%	0.0	3.3	_	-
MCM	0.79	0.05	0.63	0.04	0.03	0.85	100%	0.0	13.2	_	_
MCM with efficiency augmentation	0.36	0.02	0.31	0.02	0.02	0.59	98%	0.0	8.8	_	-
R-learning	0.36	0.02	0.31	0.02	0.02	0.59	99%	-0.1	9.1	_	-

Table D.8: Performance measure for ITE1 with random assignment and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Infeasible	1.34	0.00	0.92	0.98	0.00	0.14	59%	0.0	3.1	0.76	0.48
Conditional mean regression	5.54	0.03	4.64	1.21	0.05	1.81	6%	0.0	3.0	0.18	0.93
MOM IPW	6.56	0.03	5.88	1.25	0.05	2.02	10%	0.0	3.1	0.12	1.16
MOM DR	4.27	0.02	3.08	1.25	0.05	1.35	6%	0.0	3.0	0.18	0.45
Causal Forest	5.24	0.03	4.41	1.22	0.05	1.71	11%	0.0	3.1	0.17	0.79
Causal Forest with local centering	5.23	0.02	4.40	1.22	0.06	1.71	12%	0.0	3.1	0.17	0.81
Lasso:											
Infeasible	1.17	0.00	0.68	0.83	0.01	0.33	66%	0.0	3.5	0.79	0.60
Conditional mean regression	12.87	0.11	12.01	1.12	0.02	3.26	86%	0.0	4.1	0.15	3.53
MOM IPW	6.72	0.12	5.61	1.31	0.03	1.93	100%	0.0	14.8	0.06	1.06
MOM DR	4.57	0.04	3.25	1.31	0.03	1.32	100%	0.0	9.0	0.11	0.38
MCM	6.64	0.12	5.52	1.34	-0.02	1.85	100%	0.0	16.1	0.03	0.98
MCM with efficiency augmentation	4.49	0.04	3.22	1.32	0.04	1.28	100%	0.0	8.0	0.11	0.36
R-learning	4.52	0.04	3.29	1.32	0.04	1.29	100%	-0.1	8.2	0.11	0.37
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible	1.17	0.00	0.73	0.90	0.00	0.12	22%	0.0	3.0	0.79	0.54
Conditional mean regression	4.24	0.03	3.51	1.13	0.06	1.52	5%	0.0	3.0	0.29	0.86
MOM IPW	4.81	0.02	4.07	1.21	0.07	1.58	7%	0.0	3.0	0.18	0.83
MOM DR	3.20	0.01	2.18	1.21	0.05	0.99	7%	0.0	3.0	0.29	0.34
Causal Forest	3.78	0.02	2.96	1.17	0.06	1.31	7%	0.0	3.0	0.27	0.60
Causal Forest with local centering	3.78	0.02	2.99	1.16	0.06	1.32	9%	0.0	3.0	0.27	0.61
Lasso:	91,0	0.02			0.00		0,0	0.0		0	0.02
Infeasible	1.00	0.00	0.49	0.77	0.01	0.23	15%	0.0	3.1	0.82	0.66
Conditional mean regression	7.39	0.07	6.73	0.96	0.03	2.40	26%	0.0	3.3	0.32	2.31
MOM IPW	3.71	0.05	2.18	1.32	0.05	0.92	100%	-0.2	11.9	0.12	0.25
MOM DR	3.15	0.02	1.79	1.29	0.04	0.70	100%	-0.2	7.1	0.26	0.14
MCM	3.80	0.06	2.15	1.34	-0.01	0.88	100%	-0.1	14.1	0.06	0.23
MCM with efficiency augmentation	3.20	0.02	1.81	1.30	0.04	0.69	100%	-0.2	8.0	0.23	0.13
R-learning	3.19	0.02	1.81	1.30	0.04	0.69	99%	-0.2	7.9	0.24	0.13

Table D.9: Performance measure for ITE2 with random assignment and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Infeasible	18.84	0.01	12.79	3.66	-0.04	0.54	39%	0.0	3.1	0.74	0.47
Conditional mean regression	26.71	0.05	15.28	3.93	0.04	1.92	14%	0.0	3.0	0.62	0.29
MOM IPW	36.35	0.06	17.07	4.42	0.09	2.06	15%	0.0	3.1	0.38	0.13
MOM DR	31.96	0.06	13.75	4.32	0.08	1.45	19%	0.0	3.0	0.56	0.10
Causal Forest	29.83	0.07	15.28	4.13	0.06	1.92	32%	0.0	3.1	0.57	0.19
Causal Forest with local centering	29.16	0.07	15.57	4.09	0.06	1.93	38%	0.0	3.1	0.58	0.21
Lasso:											
Infeasible	16.81	0.01	10.77	3.22	-0.03	1.23	72%	0.0	3.4	0.77	0.57
Conditional mean regression	31.81	0.10	25.66	3.64	-0.02	3.54	78%	0.0	3.6	0.57	0.66
MOM IPW	41.74	0.12	20.93	4.61	0.03	2.37	100%	-0.2	12.2	0.25	0.14
MOM DR	33.63	0.18	16.60	4.30	0.02	2.00	100%	-0.2	17.0	0.51	0.15
MCM	43.79	0.12	20.62	4.76	-0.03	2.11	100%	-0.1	11.6	0.13	0.10
MCM with efficiency augmentation	34.08	0.08	16.63	4.34	0.02	1.92	100%	0.0	5.6	0.49	0.14
R-learning	34.08	0.09	16.66	4.34	0.03	1.93	100%	0.0	7.0	0.49	0.14
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible	16.68	0.01	11.34	3.43	-0.04	0.45	15%	0.0	3.0	0.78	0.52
Conditional mean regression	21.70	0.03	13.64	3.65	0.01	1.57	8%	0.0	3.0	0.70	0.51
MOM IPW	29.57	0.09	13.96	4.15	0.08	1.69	8%	0.0	3.0	0.57	0.17
MOM DR	23.51	0.05	13.50	3.91	0.03	1.16	8%	0.0	3.0	0.69	0.25
Causal Forest	22.07	0.04	14.11	3.75	0.01	1.53	8%	0.0	3.0	0.70	0.41
Causal Forest with local centering	21.75	0.04	14.24	3.72	0.00	1.54	10%	0.0	3.0	0.70	0.43
Lasso:											
Infeasible	14.49	0.01	8.33	3.00	-0.03	0.86	15%	0.0	3.1	0.81	0.63
Conditional mean regression	22.63	0.06	16.83	3.30	-0.02	2.56	23%	0.0	3.2	0.70	0.70
MOM IPW	30.93	0.14	15.01	4.22	0.02	1.70	99%	0.0	5.4	0.56	0.15
MOM DR	23.61	0.06	13.54	3.86	-0.01	1.38	85%	0.0	4.1	0.69	0.27
MCM	36.54	0.14	15.71	4.56	-0.03	1.36	100%	0.0	7.7	0.43	0.06
MCM with efficiency augmentation	24.21	0.06	13.76	3.91	0.00	1.36	89%	-0.1	4.4	0.68	0.26
R-learning	24.21	0.06	13.76	3.91	0.00	1.36	89%	0.0	4.4	0.68	0.26

D.1.4 ITE with random assignment and random noise

The relative performances for the ITEs with noise but randomized treatment assignment are very close to their selective equivalents. The mean SDs in the selective and randomized settings are very similar. Also within the randomized setting, the mean absolute biases are nearly identical for all estimators. The differences between selective and randomized settings are thus only driven by different capabilities of the estimators to correct for selection bias.

Table D.10: Performance measure for ITE0 with random assignment and random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Infeasible	0.79	0.00	0.82	0.62	0.00	0.05	83%	-0.2	3.2	0.04	0.00
Conditional mean regression	4.04	0.03	3.75	0.62	0.02	1.80	6%	0.0	3.0	0.00	3.36
MOM IPW	4.89	0.03	4.51	0.62	0.01	2.01	10%	0.0	3.1	0.00	4.40
MOM DR	2.64	0.02	2.43	0.62	0.01	1.36	6%	0.0	3.0	0.00	1.67
Causal Forest	3.71	0.03	3.43	0.62	0.01	1.70	9%	0.0	3.1	0.00	2.90
Causal Forest with local centering	3.71	0.02	3.41	0.62	0.01	1.70	10%	0.0	3.1	0.00	2.97
Lasso:											
Infeasible	0.79	0.00	0.86	0.62	0.00	0.05	100%	-1.3	21.4	0.02	0.00
Conditional mean regression	11.07	0.10	10.13	0.63	0.01	3.14	87%	0.0	4.0	0.00	12.32
MOM IPW	4.54	0.12	3.92	0.62	0.01	1.89	100%	0.1	15.4	0.00	3.94
MOM DR	2.45	0.05	2.17	0.62	0.02	1.27	100%	0.1	9.7	0.00	1.33
MCM	4.50	0.11	3.86	0.61	-0.02	1.86	100%	0.1	14.3	0.00	3.88
MCM with efficiency augmentation	2.35	0.04	2.07	0.62	0.02	1.23	100%	0.0	8.3	0.00	1.22
R-learning	2.32	0.04	2.03	0.62	0.02	1.22	100%	0.1	8.1	0.00	1.18
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible	0.78	0.00	0.78	0.62	0.00	0.04	30%	-0.2	3.1	0.06	0.00
Conditional mean regression	3.08	0.03	2.76	0.63	0.03	1.50	5%	0.0	3.0	0.00	2.76
MOM IPW	3.30	0.02	2.91	0.63	0.03	1.57	7%	0.0	3.0	0.00	3.03
MOM DR	1.75	0.01	1.54	0.63	0.02	0.98	6%	0.0	3.0	0.00	1.09
Causal Forest	2.47	0.01	2.19	0.63	0.02	1.29	7%	0.0	3.0	0.00	1.97
Causal Forest with local centering	2.49	0.01	2.22	0.63	0.02	1.30	7%	0.0	3.0	0.00	2.01
Lasso:											
Infeasible	0.78	0.00	0.82	0.62	0.00	0.04	100%	-0.1	7.1	0.04	0.00
Conditional mean regression	6.33	0.07	5.78	0.64	0.02	2.29	33%	0.0	3.4	0.00	6.90
MOM IPW	1.57	0.05	1.30	0.63	0.03	0.86	100%	-0.1	13.2	0.00	0.81
MOM DR	1.15	0.02	1.05	0.62	0.02	0.59	98%	0.1	8.2	0.00	0.29
MCM	1.55	0.05	1.30	0.62	0.00	0.84	100%	0.0	14.2	0.00	0.79
MCM with efficiency augmentation	1.13	0.02	1.03	0.62	0.02	0.58	97%	0.0	9.0	0.00	0.27
R-learning	1.13	0.02	1.03	0.62	0.02	0.57	96%	-0.1	8.7	0.00	0.27

Table D.11: Performance measure for ITE1 with random assignment and random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Infeasible	2.98	0.00	1.03	1.29	0.01	0.15	71%	-0.2	3.2	0.23	0.03
Conditional mean regression	6.28	0.03	4.22	1.28	0.04	1.79	7%	0.0	3.0	0.02	0.84
MOM IPW	7.16	0.03	5.50	1.27	0.03	2.01	10%	0.0	3.0	0.01	1.11
MOM DR	4.88	0.02	2.76	1.27	0.03	1.35	6%	0.0	3.0	0.02	0.42
Causal Forest	5.94	0.03	3.91	1.27	0.03	1.70	10%	0.0	3.1	0.02	0.73
Causal Forest with local centering	5.93	0.02	3.92	1.27	0.03	1.70	12%	0.0	3.1	0.02	0.74
Lasso:											
Infeasible	3.00	0.00	1.08	1.28	0.01	0.21	100%	-0.5	7.7	0.21	0.04
Conditional mean regression	13.26	0.10	11.73	1.28	0.03	3.14	87%	0.0	4.0	0.02	3.10
MOM IPW	6.99	0.13	5.34	1.28	0.02	1.92	100%	-0.2	17.2	0.01	1.04
MOM DR	4.78	0.05	2.74	1.28	0.03	1.28	100%	0.1	9.8	0.01	0.34
MCM	6.78	0.11	5.27	1.27	-0.04	1.85	100%	0.1	15.1	0.00	0.96
MCM with efficiency augmentation	4.69	0.04	2.67	1.28	0.03	1.24	100%	0.0	8.5	0.01	0.32
R-learning	4.65	0.04	2.64	1.28	0.03	1.23	100%	0.0	8.4	0.01	0.31
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible	2.93	0.00	1.27	1.30	0.01	0.11	26%	-0.1	3.1	0.25	0.06
Conditional mean regression	5.31	0.03	3.33	1.28	0.05	1.50	5%	0.0	3.0	0.03	0.70
MOM IPW	5.57	0.02	3.74	1.28	0.04	1.57	6%	0.0	3.0	0.02	0.77
MOM DR	4.00	0.01	1.90	1.28	0.04	0.98	6%	0.0	3.0	0.04	0.28
Causal Forest	4.70	0.01	2.67	1.28	0.04	1.29	7%	0.0	3.0	0.03	0.50
Causal Forest with local centering	4.71	0.01	2.69	1.28	0.04	1.29	8%	0.0	3.0	0.03	0.51
Lasso:											
Infeasible	2.93	0.00	1.08	1.28	0.01	0.16	82%	-0.1	4.0	0.25	0.06
Conditional mean regression	8.67	0.08	7.30	1.28	0.03	2.33	32%	0.0	3.3	0.04	1.83
MOM IPW	3.89	0.05	1.86	1.28	0.04	0.86	100%	0.0	11.7	0.01	0.21
MOM DR	3.46	0.02	1.41	1.28	0.04	0.59	98%	-0.1	8.8	0.02	0.07
MCM	3.92	0.05	1.86	1.27	-0.02	0.86	100%	0.0	13.8	0.01	0.21
MCM with efficiency augmentation	3.45	0.02	1.44	1.28	0.03	0.58	99%	0.0	9.0	0.02	0.07
R-learning	3.46	0.02	1.43	1.28	0.03	0.59	98%	-0.1	8.7	0.02	0.07

Table D.12: Performance measure for ITE2 with random assignment and random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Infeasible	38.46	0.01	8.96	4.43	0.09	0.52	66%	-0.2	3.2	0.19	0.02
Conditional mean regression	41.88	0.03	13.82	4.44	0.12	1.76	6%	0.0	3.0	0.06	0.07
MOM IPW	43.04	0.03	14.92	4.44	0.12	2.01	11%	0.0	3.0	0.04	0.09
MOM DR	40.76	0.02	12.86	4.44	0.10	1.35	5%	0.0	3.0	0.06	0.04
Causal Forest	41.67	0.03	13.85	4.44	0.11	1.69	9%	0.0	3.0	0.06	0.06
Causal Forest with local centering	41.68	0.02	13.84	4.44	0.12	1.70	10%	0.0	3.0	0.06	0.06
Lasso:											
Infeasible	38.66	0.01	8.88	4.42	0.08	0.71	100%	-0.6	8.3	0.18	0.03
Conditional mean regression	48.39	0.10	21.71	4.43	0.11	3.13	89%	0.0	3.9	0.05	0.26
MOM IPW	43.34	0.13	16.17	4.46	0.10	1.94	100%	-0.2	16.9	0.02	0.08
MOM DR	41.11	0.04	13.50	4.46	0.11	1.32	100%	0.0	7.9	0.04	0.03
MCM	43.26	0.12	16.45	4.47	-0.08	1.87	100%	0.0	14.3	0.01	0.08
MCM with efficiency augmentation	41.09	0.04	13.51	4.46	0.11	1.30	100%	0.0	8.2	0.04	0.03
R-learning	41.09	0.04	13.50	4.46	0.11	1.30	100%	0.0	8.2	0.04	0.03
				400	00 obse	rvatio	ns				
Random Forest:											
Infeasible	37.86	0.00	6.30	4.43	0.07	0.41	31%	-0.2	3.1	0.22	0.04
Conditional mean regression	40.51	0.03	11.74	4.43	0.13	1.47	5%	0.0	3.0	0.10	0.07
MOM IPW	41.23	0.02	13.48	4.43	0.12	1.57	6%	0.0	3.0	0.06	0.06
MOM DR	39.63	0.02	11.70	4.43	0.10	0.99	6%	0.0	3.0	0.10	0.03
Causal Forest	40.15	0.02	11.72	4.43	0.12	1.30	7%	0.0	3.0	0.09	0.05
Causal Forest with local centering	40.17	0.02	11.68	4.43	0.11	1.32	8%	0.0	3.0	0.09	0.05
Lasso:											
Infeasible	37.84	0.00	6.69	4.40	0.07	0.53	84%	-0.1	4.2	0.22	0.04
Conditional mean regression	43.27	0.07	14.64	4.42	0.09	2.33	27%	0.0	3.3	0.11	0.18
MOM IPW	40.27	0.05	12.33	4.46	0.12	0.94	100%	-0.3	11.6	0.05	0.02
MOM DR	39.67	0.02	11.83	4.45	0.11	0.73	99%	-0.1	6.8	0.09	0.01
MCM	40.38	0.05	11.46	4.47	-0.06	0.87	100%	-0.3	13.4	0.02	0.02
MCM with efficiency augmentation	39.74	0.02	11.81	4.45	0.11	0.73	100%	0.0	7.8	0.08	0.01
R-learning	39.73	0.02	11.86	4.45	0.11	0.73	100%	0.0	7.7	0.08	0.01

D.1.5 ITE without censoring

This appendix shows the results for an alternative DGP that ignores the natural bounds of our outcome variable. It takes the following form similar to Equations 30 to 21:

$$\omega(x) = \sin\left(1.25\pi \frac{p^{HLM}(x)}{\max(p^{HLM}(x))}\right),\tag{30}$$

$$\Omega(x) = \alpha \frac{\omega(x) - \bar{\omega}}{SD(\omega(x))} + \varepsilon_i, \tag{31}$$

$$\xi_{nc}(x) = \begin{cases} \lfloor \Omega(x) & \text{if } \Omega(x) - \lfloor \Omega(x) < u_i \\ \Omega(x) \end{pmatrix} & \text{if } \Omega(x) - \lfloor \Omega(x) \ge u_i \end{cases}$$
(32)

where u_i is uniformly distributed between zero and one. It is similar to the baseline ITE1 without the censoring.

We run this robustness check for two purposes. First, to investigate whether our results are sensitive to such modifications of the DGP. Second, we know the true IATE in this formulation because $\xi_{nc}(x)$ does not depend on the non-treated outcome Y_i^0 via censoring.

Table D.13 shows the performance measure as in all tables above comparing the estimated IATEs to the *true ITEs* in the validation sample. The only striking difference to the other results is the very high mean MSE of Lasso MOM IPW. This is driven by extreme outliers as indicated by the median MSE that is comparable to other methods.

Table D.14 shows the performance measures when comparing the estimated IATEs to the *true IATEs*. The only differences compared to Table D.13 are the lower MSE measures that are driven by lower mean absolute biases. Although the level of these measures is changed, the ordering remains the same. This is expected because we get rid of the irreducible noise component that enters as bias if the true ITEs is considered as benchmark. This exercise illustrates that the performance measures based on the true ITEs as benchmark lead to the same conclusions as if we would know the true IATE.

Table D.13: Performance measure for ITE1 without censoring

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	0 obser	vation	ıs				
Random Forest:											
Infeasible	2.17	0.00	0.92	1.15	0.00	0.19	79%	-0.1	3.2	0.77	0.49
Conditional mean regression	9.50	0.06	5.98	1.92	1.34	1.86	7%	0.0	3.0	0.18	0.66
MOM IPW	10.83	0.08	7.73	1.85	1.31	2.29	17%	0.0	3.1	0.62	2.60
MOM DR	6.75	0.03	4.28	1.78	0.88	1.35	12%	0.0	3.0	0.22	0.29
Causal Forest	9.57	0.06	5.79	2.00	1.55	1.75	14%	0.0	3.1	0.19	0.54
Causal Forest with local centering	8.21	0.04	5.89	1.76	0.84	1.84	15%	0.0	3.1	0.16	0.56
Lasso:											
Infeasible	1.86	0.00	0.90	1.01	0.00	0.42	75%	-0.1	3.5	0.80	0.61
Conditional mean regression	16.87	0.12	14.74	1.68	1.12	3.43	90%	-0.1	4.0	0.20	2.50
MOM IPW	1048.99	1034.37	12.59	1.41	1.10	5.09	100%	1.0	41.5	0.41	200.92
MOM DR	8.83	0.80	6.12	1.86	0.80	1.80	100%	-8.4	223.4	0.13	0.38
MCM	13.67	0.21	11.87	1.27	0.68	3.17	100%	-0.1	6.9	0.42	2.68
MCM with efficiency augmentation	7.35	0.05	5.17	1.85	0.79	1.43	100%	-0.2	7.8	0.15	0.28
R-learning	7.59	0.06	5.12	1.89	0.90	1.39	100%	-0.1	10.6	0.11	0.26
				400	0 obser	vation	ıs				
Random Forest:											
Infeasible	1.91	0.00	0.86	1.07	0.00	0.16	29%	-0.1	3.1	0.80	0.55
Conditional mean regression	7.80	0.06	4.77	1.83	1.20	1.56	5%	0.0	3.0	0.26	0.61
MOM IPW	6.89	0.05	4.87	1.42	0.91	1.84	9%	0.0	3.0	0.59	1.66
MOM DR	5.12	0.03	3.28	1.63	0.58	1.05	15%	0.0	3.1	0.35	0.25
Causal Forest	7.77	0.06	4.26	1.91	1.47	1.36	8%	0.0	3.0	0.29	0.44
Causal Forest with local centering	6.14	0.03	4.24	1.62	0.64	1.44	11%	0.0	3.1	0.28	0.44
Lasso:											
Infeasible	1.59	0.00	0.73	0.95	0.00	0.28	16%	0.0	3.1	0.83	0.67
Conditional mean regression	9.93	0.09	8.44	1.38	0.78	2.53	35%	-0.1	3.3	0.38	1.79
MOM IPW	7.63	0.14	6.27	1.19	0.74	2.19	97%	0.1	5.4	0.53	1.61
MOM DR	5.34	0.04	3.08	1.76	0.51	0.85	99%	-0.4	9.3	0.31	0.14
MCM	7.34	0.14	5.80	1.20	0.49	2.09	100%	-0.1	5.4	0.49	1.44
MCM with efficiency augmentation	5.28	0.03	2.95	1.77	0.47	0.82	99%	-0.2	7.6	0.30	0.13
R-learning	5.49	0.03	3.22	1.81	0.60	0.79	99%	-0.2	7.5	0.27	0.11

Table D.14: Performance measure for ITE1 without censoring with true IATE as benchmark

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	0 obser	vation	ıs				
Random Forest:											
Infeasible	0.98	0.00	0.48	0.77	0.00	0.19	79%	-0.1	3.2	0.77	0.49
Conditional mean regression	8.31	0.06	5.26	1.69	1.34	1.86	7%	0.0	3.0	0.18	0.66
MOM IPW	9.64	0.08	7.19	1.66	1.31	2.29	17%	0.0	3.1	0.62	2.60
MOM DR	5.56	0.03	3.42	1.56	0.88	1.35	12%	0.0	3.0	0.22	0.29
Causal Forest	8.37	0.06	4.96	1.77	1.55	1.75	14%	0.0	3.1	0.19	0.54
Causal Forest with local centering	7.01	0.04	5.09	1.54	0.84	1.84	15%	0.0	3.1	0.16	0.56
Lasso:											
Infeasible	0.68	0.00	0.41	0.53	0.00	0.42	75%	-0.1	3.5	0.80	0.61
Conditional mean regression	15.68	0.12	14.08	1.44	1.12	3.43	90%	-0.1	4.0	0.20	2.50
MOM IPW	1047.83	1034.38	11.76	1.19	1.10	5.09	100%	1.0	41.5	0.41	200.92
MOM DR	7.64	0.80	4.77	1.68	0.80	1.80	100%	-8.4	223.4	0.13	0.38
MCM	12.49	0.21	10.94	0.95	0.68	3.17	100%	-0.1	6.9	0.42	2.68
MCM with efficiency augmentation	6.16	0.05	3.91	1.66	0.79	1.43	100%	-0.2	7.8	0.15	0.28
R-learning	6.40	0.06	3.72	1.71	0.90	1.39	100%	-0.1	10.6	0.11	0.26
				400	0 obser	vation	ıs				
Random Forest:											
Infeasible	0.72	0.00	0.33	0.64	0.00	0.16	29%	-0.1	3.1	0.80	0.55
Conditional mean regression	6.60	0.06	4.18	1.59	1.20	1.56	5%	0.0	3.0	0.27	0.61
MOM IPW	5.72	0.05	4.34	1.15	0.91	1.84	9%	0.0	3.0	0.59	1.66
MOM DR	3.92	0.03	2.76	1.41	0.58	1.05	15%	0.0	3.1	0.35	0.25
Causal Forest	6.57	0.06	3.57	1.67	1.47	1.36	8%	0.0	3.0	0.29	0.44
Causal Forest with local centering	4.94	0.03	3.68	1.39	0.64	1.44	11%	0.0	3.1	0.28	0.44
Lasso:											
Infeasible	0.42	0.00	0.20	0.39	0.00	0.28	16%	0.0	3.1	0.83	0.67
Conditional mean regression	8.75	0.09	7.84	1.10	0.78	2.53	35%	-0.1	3.3	0.38	1.79
MOM IPW	6.47	0.14	5.63	0.88	0.74	2.19	97%	0.1	5.4	0.52	1.61
MOM DR	4.15	0.04	2.94	1.56	0.50	0.85	99%	-0.4	9.3	0.31	0.14
MCM	6.17	0.14	4.92	0.87	0.49	2.09	100%	-0.1	5.4	0.49	1.44
MCM with efficiency augmentation	4.09	0.03	3.00	1.57	0.47	0.82	99%	-0.2	7.6	0.30	0.13
R-learning	4.31	0.03	2.87	1.61	0.60	0.79	99%	-0.2	7.5	0.27	0.11

D.2 Results for GATE estimation

The Appendices D.2.1 to D.2.4 show the full results for GATE estimation in the 24 DGP-sample size combinations. The performance measures are the same as for IATE estimation. However, we omit the infeasible benchmark because it is not clear how it should be constructed for GATEs.

D.2.1 GATEs from ITE with selection and without random noise

The additional performance measures for the baseline ITE0 do not change the conclusions in the main text. The mean bias shows that the substantial positive biases remain due to selection bias. This is in line with the IATE results and shows that averaging the IATEs does not remove the selection bias.

Also the GATE estimation of ITE1 and ITE2 without random noise shows patterns that are already observed for their baseline equivalents with noise. We observe that some estimators that show low mean absolute bias but high mean SD for the IATEs become competitive by averaging out the noise. In particular, Random Forest MOM IPW shows very low mean MSE for ITE1 with 4,000 observations and ITE2 with 1,000 observations. Similarly, Lasso mean regression is locally very successful for ITE2 in the 4,000 observations sample. While such competitive performances are usually not consistent for mean MSE and median MSE in the case of IATE estimation, they are confirmed for GATE estimation. However, we observe no systematic pattern that explains under which circumstances which noisy IATE estimator provides a good GATE estimator. Thus, the four consistently well performing IATE estimators seem to be also the dominant choice for the GATE estimation.

Table D.15: Performance measures for GATE of ITE0 with selection and without random noise (baseline)

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	1.77	0.03	1.32	0.55	0.55	1.19	22%	0.0	3.1	-	-
MOM IPW	4.44	0.05	2.12	1.59	0.05	1.16	47%	-0.1	3.2	-	-
MOM DR	0.87	0.02	0.93	0.38	0.38	0.85	20%	0.0	3.1	-	-
Causal Forest	1.44	0.03	1.22	0.70	0.70	0.96	17%	0.0	3.1	-	-
Causal Forest with local centering	1.08	0.02	1.03	0.33	0.33	0.99	8%	0.0	3.0	-	-
Lasso:											
Conditional mean regression	3.35	0.04	1.77	0.55	0.54	1.70	34%	0.0	3.1	-	-
MOM IPW	3.15	0.07	1.72	0.78	0.32	1.50	100%	-0.1	4.6	-	-
MOM DR	38.85	37.73	6.13	0.59	0.59	6.20	100%	43.1	1901.5	-	-
MCM	4.56	0.10	1.93	1.20	0.03	1.62	100%	-0.3	4.0	-	-
MCM with efficiency augmentation	1.04	0.03	1.02	0.41	0.41	0.93	66%	-0.1	3.5	_	-
R-learning	1.04	0.03	1.01	0.44	0.44	0.92	73%	-0.1	3.7	-	-
				400	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	1.07	0.03	1.01	0.49	0.47	0.86	8%	0.0	3.1	-	-
MOM IPW	1.12	0.02	0.92	0.67	0.21	0.66	6%	0.0	3.0	-	-
MOM DR	0.30	0.01	0.55	0.25	0.25	0.48	14%	0.0	3.1	-	-
Causal Forest	0.74	0.03	0.86	0.64	0.64	0.53	0%	0.0	2.9	-	-
Causal Forest with local centering	0.35	0.01	0.59	0.22	0.22	0.54	3%	0.0	3.0	-	-
Lasso:											
Conditional mean regression	1.45	0.03	1.16	0.47	0.42	1.06	6%	0.0	3.1	-	-
MOM IPW	1.19	0.04	1.06	0.53	0.26	0.87	86%	-0.1	3.7	-	-
MOM DR	0.30	0.02	0.54	0.31	0.31	0.45	38%	0.0	3.4	-	-
MCM	1.65	0.06	1.14	0.75	0.07	0.92	94%	-0.1	3.8	-	-
MCM with efficiency augmentation	0.27	0.01	0.51	0.26	0.26	0.45	55%	-0.1	3.7	-	-
R-learning	0.27	0.02	0.52	0.28	0.28	0.44	25%	0.1	3.4		

Table D.16: Performance measures for GATE of ITE1 with selection and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	ervatio	ns				
Random Forest:											
Conditional mean regression	4.39	0.06	1.84	1.41	1.40	1.20	22%	0.0	3.1	-0.11	0.33
MOM IPW	3.19	0.04	1.55	1.10	0.72	1.18	44%	-0.1	3.2	0.56	1.85
MOM DR	2.66	0.04	1.30	1.09	1.04	0.83	22%	-0.1	3.1	-0.06	0.09
Causal Forest	4.40	0.06	1.85	1.62	1.62	0.96	20%	0.0	3.1	-0.22	0.14
Causal Forest with local centering	2.87	0.04	1.36	1.08	1.01	0.98	14%	0.0	3.0	-0.03	0.12
Lasso:											
Conditional mean regression	5.11	0.06	2.24	1.17	1.15	1.76	42%	-0.1	3.1	0.07	0.95
MOM IPW	3.50	0.06	1.83	0.89	0.89	1.60	100%	-0.1	4.2	0.32	1.15
MOM DR	37.54	34.31	6.07	1.26	1.21	5.91	100%	42.8	1887.4	-1.16	0.09
MCM	3.43	0.07	1.71	0.58	0.45	1.68	100%	-0.3	3.9	0.43	1.52
MCM with efficiency augmentation	3.06	0.04	1.42	1.16	1.03	0.95	78%	-0.1	3.8	-1.42	0.08
R-learning	3.34	0.05	1.47	1.23	1.14	0.93	78%	-0.1	3.8	-1.62	0.06
				400	00 obse	ervatio	ns				
Random Forest:											
Conditional mean regression	3.15	0.05	1.46	1.26	1.23	0.88	6%	0.0	3.1	0.03	0.31
MOM IPW	1.09	0.03	0.93	0.69	0.69	0.68	16%	0.0	3.1	0.52	0.65
MOM DR	1.47	0.03	0.90	0.87	0.75	0.48	25%	0.0	3.2	0.26	0.08
Causal Forest	3.27	0.06	1.62	1.51	1.51	0.54	3%	0.0	3.0	-0.05	0.12
Causal Forest with local centering	1.67	0.03	0.96	0.91	0.80	0.55	8%	0.0	3.1	0.19	0.09
Lasso:											
Conditional mean regression	2.08	0.04	1.38	0.77	0.76	1.11	8%	0.0	3.1	0.34	0.67
MOM IPW	1.43	0.04	1.20	0.65	0.65	0.96	64%	0.0	3.5	0.46	0.74
MOM DR	1.90	0.04	0.89	1.03	0.79	0.50	69%	-0.1	3.5	-1.41	0.04
MCM	1.28	0.04	1.00	0.42	0.40	0.97	67%	-0.1	3.4	0.47	0.72
MCM with efficiency augmentation	1.81	0.04	0.89	1.01	0.75	0.49	69%	-0.1	3.4	-0.77	0.04
R-learning	2.09	0.04	0.94	1.09	0.87	0.48	66%	-0.1	3.8	-1.29	0.03

Table D.17: Performance measures for GATE of ITE2 with selection and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	ervatio	$\mathbf{n}\mathbf{s}$				
Random Forest:											
Conditional mean regression	19.52	0.14	3.50	3.30	3.05	1.29	28%	0.0	3.1	0.28	0.08
MOM IPW	8.38	0.09	2.54	2.30	2.30	1.24	45%	-0.1	3.1	0.56	0.34
MOM DR	19.64	0.09	2.71	3.44	2.68	0.85	41%	-0.1	3.1	0.24	0.03
Causal Forest	27.12	0.15	3.87	3.97	3.84	1.02	44%	-0.1	3.2	0.00	0.04
Causal Forest with local centering	17.59	0.12	2.75	3.20	2.54	1.06	47%	-0.1	3.2	0.32	0.05
Lasso:											
Conditional mean regression	9.28	0.10	2.51	1.88	1.77	1.88	30%	0.0	3.1	0.52	0.28
MOM IPW	9.81	0.12	2.77	2.14	2.11	1.84	97%	-0.1	3.7	0.51	0.29
MOM DR	51.94	34.59	6.64	3.25	2.35	5.97	100%	41.5	1810.7	0.11	0.05
MCM	9.29	0.12	2.35	1.95	1.53	1.80	94%	-0.2	3.6	0.52	0.21
MCM with efficiency augmentation	17.87	0.13	2.93	3.27	2.34	1.18	94%	-0.2	3.6	0.12	0.04
R-learning	20.45	0.14	2.86	3.50	2.63	1.14	84%	-0.2	3.6	-0.20	0.03
				400	00 obse	ervatio	ns				
Random Forest:											
Conditional mean regression	8.44	0.12	2.91	2.26	2.10	0.94	19%	0.0	3.1	0.55	0.20
MOM IPW	5.23	0.08	1.62	1.78	1.68	0.73	17%	-0.1	3.1	0.60	0.25
MOM DR	7.77	0.09	1.92	2.18	1.60	0.58	2%	0.0	3.0	0.59	0.13
Causal Forest	16.66	0.19	3.38	3.21	3.16	0.72	16%	0.0	2.9	0.40	0.10
Causal Forest with local centering	5.36	0.09	1.55	1.76	1.41	0.71	3%	0.1	3.0	0.62	0.20
Lasso:											
Conditional mean regression	2.64	0.05	1.48	0.96	0.89	1.11	5%	0.0	3.0	0.65	0.40
MOM IPW	3.51	0.07	1.72	1.35	1.33	1.05	25%	0.0	3.2	0.63	0.36
MOM DR	5.50	0.09	1.81	1.85	1.19	0.71	27%	0.0	3.2	0.64	0.16
MCM	6.41	0.13	1.88	1.85	1.21	1.08	48%	-0.1	3.4	0.61	0.17
MCM with efficiency augmentation	6.60	0.09	2.02	2.05	1.35	0.69	36%	0.0	3.2	0.62	0.14
R-learning	7.86	0.11	2.17	2.23	1.55	0.72	41%	0.1	3.2	0.60	0.12

D.2.2 GATEs from ITE with selection and random noise

The GATE results for ITE0 with random noise are very similar to the baseline without noise. Besides Lasso MOM DR that performs even worse, all estimators show very similar performance with differences only at the second digit of most performance measures. This shows how the additional noise is averaged out on the group level.

The additional performance measures for the baseline ITE1 and ITE2 confirm the results that are discussed in the main text.

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Table D.18: Performance measures for GATE of ITE0 with selection and random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				10	00 obse	ervatio	ns				
Random Forest:											
Conditional mean regression	1.75	0.03	1.30	0.53	0.53	1.18	25%	0.0	3.1	0.01	1.09
MOM IPW	4.57	0.05	2.09	1.61	0.02	1.16	41%	-0.1	3.2	-0.01	4.95
MOM DR	0.84	0.02	0.92	0.35	0.35	0.84	20%	0.0	3.1	0.02	0.26
Causal Forest	1.41	0.03	1.18	0.68	0.68	0.95	16%	0.0	3.1	0.03	0.40
Causal Forest with local centering $Lasso:$	1.06	0.02	1.02	0.30	0.30	0.98	20%	0.0	3.1	0.02	0.41
Conditional mean regression	3.28	0.04	1.74	0.54	0.52	1.69	33%	0.0	3.1	0.01	3.00
MOM IPW	3.14	0.07	1.71	0.80	0.29	1.49	100%	-0.1	4.5	0.02	2.90
MOM DR	72.02	70.92	6.06	0.60	0.60	7.64	100%	43.0	1900.7	0.13	17.68
MCM	4.55	0.09	1.91	1.20	-0.03	1.61	98%	-0.3	3.9	0.01	4.74
MCM with efficiency augmentation	1.03	0.03	1.01	0.39	0.39	0.93	75%	0.0	3.6	0.09	0.22
R-learning	1.03	0.03	1.01	0.42	0.42	0.92	73%	0.0	3.7	0.11	0.21
				40	00 obse	ervatio	ns				
Random Forest:											
Conditional mean regression	1.06	0.03	1.01	0.48	0.45	0.86	6%	0.0	3.0	0.01	0.88
MOM IPW	1.16	0.02	0.96	0.69	0.19	0.66	6%	0.0	3.0	-0.01	1.17
MOM DR	0.30	0.01	0.54	0.24	0.23	0.48	13%	0.0	3.0	0.03	0.12
Causal Forest	0.72	0.03	0.85	0.62	0.62	0.53	2%	0.1	2.9	0.03	0.21
Causal Forest with local centering	0.35	0.01	0.59	0.21	0.20	0.54	5%	0.1	3.0	0.02	0.17
Lasso:											
Conditional mean regression	1.43	0.03	1.15	0.45	0.41	1.05	14%	0.0	3.1	0.01	1.39
MOM IPW	1.20	0.04	1.06	0.54	0.24	0.87	91%	0.0	3.9	0.02	1.23
MOM DR	0.30	0.02	0.54	0.29	0.29	0.45	42%	0.1	3.5	0.20	0.05
MCM	1.60	0.06	1.10	0.75	0.03	0.90	89%	-0.2	3.6	0.02	1.78
MCM with efficiency augmentation	0.27	0.01	0.51	0.24	0.24	0.45	28%	0.0	3.4	0.14	0.05
R-learning	0.28	0.01	0.52	0.26	0.26	0.44	14%	0.1	3.2	0.19	0.04

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Table D.19: Performance measures for GATE of ITE1 with selection and random noise (baseline)

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	ervatio	ns				
Random Forest:											
Conditional mean regression	2.30	0.04	1.41	0.85	0.84	1.18	20%	0.0	3.1	-0.04	0.28
MOM IPW	3.84	0.04	1.89	1.41	0.26	1.17	41%	-0.1	3.1	0.20	1.43
MOM DR	1.16	0.02	1.01	0.59	0.59	0.83	20%	0.0	3.1	-0.03	0.07
Causal Forest	2.04	0.04	1.32	1.01	1.01	0.95	19%	0.0	3.1	-0.06	0.11
Causal Forest with local centering	1.38	0.03	1.11	0.56	0.56	0.97	17%	0.0	3.1	-0.02	0.10
Lasso:											
Conditional mean regression	3.68	0.05	1.92	0.78	0.76	1.69	41%	-0.1	3.1	-0.01	0.77
MOM IPW	3.03	0.06	1.70	0.65	0.51	1.53	100%	-0.2	5.3	0.09	0.84
MOM DR	39.33	37.80	6.18	0.79	0.79	6.20	100%	43.0	1899.5	-0.30	0.08
MCM	3.98	0.09	1.79	0.93	0.11	1.65	97%	-0.2	4.1	0.15	1.32
MCM with efficiency augmentation	1.40	0.03	1.11	0.62	0.61	0.93	80%	-0.1	3.7	-0.32	0.06
R-learning	1.45	0.03	1.13	0.68	0.67	0.91	75%	0.0	3.6	-0.31	0.05
				400	00 obse	ervatio	ns				
Random Forest:											
Conditional mean regression	1.53	0.04	1.11	0.76	0.74	0.86	6%	0.0	3.1	-0.02	0.23
MOM IPW	0.99	0.02	0.81	0.54	0.37	0.67	11%	0.0	3.0	0.17	0.39
MOM DR	0.49	0.02	0.61	0.42	0.41	0.48	17%	0.0	3.1	0.02	0.04
Causal Forest	1.26	0.04	1.00	0.93	0.93	0.53	2%	0.0	3.0	-0.03	0.06
Causal Forest with local centering	0.58	0.02	0.67	0.44	0.42	0.54	5%	0.1	3.0	0.01	0.05
Lasso:											
Conditional mean regression	1.66	0.03	1.27	0.60	0.57	1.06	14%	0.0	3.1	0.03	0.39
MOM IPW	1.17	0.04	1.04	0.47	0.38	0.89	77%	0.0	3.7	0.13	0.41
MOM DR	0.59	0.02	0.61	0.50	0.47	0.46	42%	0.0	3.5	-0.37	0.02
MCM	1.31	0.05	1.03	0.52	0.13	0.93	91%	-0.1	3.6	0.16	0.53
MCM with efficiency augmentation	0.55	0.02	0.58	0.47	0.42	0.45	39%	0.0	3.3	-0.39	0.01
R-learning	0.61	0.02	0.61	0.51	0.48	0.45	48%	0.0	3.4	-0.60	0.01

Table D.20: Performance measures for GATE of ITE2 with selection and random noise (baseline)

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	ervatio	ns				
Random Forest:											
Conditional mean regression	5.28	0.06	1.70	1.57	1.51	1.15	20%	0.0	3.1	-0.05	0.02
MOM IPW	3.75	0.04	1.61	1.25	0.77	1.18	41%	-0.1	3.1	0.16	0.15
MOM DR	3.50	0.04	1.32	1.29	1.10	0.84	23%	-0.1	3.1	-0.02	0.01
Causal Forest	5.33	0.06	1.65	1.71	1.69	0.94	25%	0.0	3.1	-0.07	0.01
Causal Forest with local centering	3.74	0.05	1.40	1.28	1.10	0.98	11%	0.0	3.0	-0.01	0.01
Lasso:											
Conditional mean regression	5.73	0.06	2.15	1.34	1.29	1.71	42%	-0.1	3.1	0.02	0.07
MOM IPW	4.00	0.06	1.94	1.01	0.95	1.59	100%	-0.1	4.1	0.09	0.09
MOM DR	30.36	26.20	4.28	1.43	1.24	4.71	100%	40.2	1737.8	-0.47	0.13
MCM	3.65	0.06	1.84	0.66	0.37	1.67	100%	-0.3	4.0	0.13	0.12
MCM with efficiency augmentation	3.94	0.05	1.42	1.35	1.10	0.95	72%	-0.1	3.7	-0.39	0.01
R-learning	4.27	0.05	1.46	1.43	1.22	0.93	72%	-0.1	3.7	-0.45	0.01
				400	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	3.97	0.06	1.45	1.44	1.36	0.85	22%	0.0	3.2	0.01	0.02
MOM IPW	1.72	0.03	1.14	0.95	0.76	0.68	14%	-0.1	3.0	0.15	0.05
MOM DR	2.19	0.03	0.95	1.08	0.81	0.49	23%	0.0	3.1	0.08	0.01
Causal Forest	4.17	0.06	1.47	1.60	1.59	0.53	8%	0.0	3.0	-0.02	0.01
Causal Forest with local centering	2.43	0.04	1.00	1.12	0.90	0.55	9%	0.0	3.0	0.06	0.01
Lasso:											
Conditional mean regression	2.59	0.05	1.39	0.95	0.89	1.09	11%	0.0	3.0	0.11	0.05
MOM IPW	1.92	0.04	1.33	0.82	0.71	0.96	45%	0.1	3.4	0.14	0.06
MOM DR	2.71	0.04	1.02	1.23	0.85	0.52	69%	-0.1	3.6	-0.33	0.00
MCM	1.58	0.04	1.15	0.61	0.33	0.97	81%	-0.1	3.6	0.16	0.06
MCM with efficiency augmentation	2.63	0.04	1.04	1.22	0.81	0.51	67%	-0.1	3.4	-0.24	0.00
R-learning	2.95	0.04	1.07	1.29	0.94	0.50	50%	-0.1	3.5	-0.33	0.00

D.2.3 GATEs from ITE with random assignment and without random noise

The GATE estimation with random assignment and without noise shows similar patterns as for the IATE estimation discussed in D.1.3. However, the outstanding performance of Lasso conditional mean regression for the large ITE2 is even more pronounced with mean absolute biases of less than halve of the next best estimator. Furthermore, both versions of Causal Forest perform best for ITE1 with 4,000 observations.

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Table D.21: Performance measures for GATE of ITE0 with random assignment and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Conditional mean regression	1.43	0.03	1.20	0.02	0.01	1.19	23%	0.0	3.1	-	-
MOM IPW	1.13	0.02	1.05	0.01	0.01	1.06	28%	0.0	3.1	-	-
MOM DR	0.75	0.02	0.87	0.01	0.01	0.87	0%	0.0	3.0	-	-
Causal Forest	0.94	0.02	0.97	0.01	0.01	0.97	11%	0.0	3.1	-	_
Causal Forest with local centering	0.90	0.02	0.94	0.02	0.01	0.94	2%	0.0	3.0	_	-
Lasso:											
Conditional mean regression	2.90	0.04	1.64	0.02	0.01	1.68	20%	0.0	3.1	-	-
MOM IPW	1.09	0.03	1.00	0.01	0.01	1.04	100%	0.0	4.3	-	_
MOM DR	0.78	0.02	0.88	0.01	0.01	0.88	61%	0.0	3.5	-	-
MCM	1.04	0.03	0.96	0.01	0.01	1.01	100%	0.0	4.3	-	_
MCM with efficiency augmentation	0.74	0.02	0.85	0.02	0.02	0.86	64%	0.0	3.4	-	_
R-learning	0.73	0.02	0.85	0.02	0.02	0.86	56%	0.0	3.4	-	-
				400	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	0.76	0.02	0.84	0.04	0.02	0.86	6%	0.0	3.1	-	-
MOM IPW	0.33	0.01	0.57	0.03	0.03	0.57	11%	0.0	3.2	-	-
MOM DR	0.20	0.01	0.45	0.02	0.02	0.45	2%	0.0	3.0	-	-
Causal Forest	0.26	0.01	0.52	0.03	0.03	0.51	0%	0.1	3.0	-	-
Causal Forest with local centering	0.25	0.01	0.50	0.02	0.02	0.50	3%	0.1	3.0	_	-
Lasso:											
Conditional mean regression	1.13	0.02	1.03	0.05	0.01	1.05	6%	0.0	3.0	-	-
MOM IPW	0.24	0.01	0.46	0.03	0.03	0.48	59%	0.1	4.7	-	-
MOM DR	0.17	0.01	0.41	0.02	0.02	0.41	6%	0.1	3.0	-	-
MCM	0.23	0.01	0.45	0.03	0.03	0.47	56%	0.2	4.4	-	-
MCM with efficiency augmentation	0.16	0.01	0.40	0.02	0.02	0.40	14%	0.1	3.1	-	-
R-learning	0.16	0.01	0.40	0.02	0.02	0.40	17%	0.1	3.1	-	-

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Table D.22: Performance measures for GATE of ITE1 with random assignment and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	2.27	0.03	1.42	0.79	0.31	1.19	17%	0.0	3.1	0.24	0.30
MOM IPW	2.12	0.03	1.33	0.85	0.34	1.05	22%	0.0	3.1	0.19	0.18
MOM DR	1.74	0.02	1.19	0.85	0.34	0.86	2%	0.0	3.0	0.28	0.09
Causal Forest	1.84	0.03	1.24	0.80	0.32	0.97	9%	0.0	3.1	0.30	0.13
Causal Forest with local centering	1.78	0.02	1.23	0.80	0.33	0.94	6%	0.0	3.0	0.29	0.13
Lasso:	0.45	0.04	1.04	0.55	0.00	1 71	1.007	0.0	0.1	0.00	0.00
Conditional mean regression	3.47	0.04	1.84	0.57	0.20	1.71	16%	0.0	3.1	0.22	0.86
MOM IPW	2.39	0.03	1.38	0.97	0.38	1.04	100%	-0.1	4.3	0.02	0.15
MOM DR	2.09	0.03	1.27	0.97	0.38	0.90	77%	-0.1	3.5	0.06	0.07
MCM	2.43	0.03	1.41	1.02	0.34	1.02	100%	0.0	4.4	0.02	0.13
MCM with efficiency augmentation	2.05	0.03	1.25	0.97	0.39	0.87	67%	0.0	3.5	0.06	0.06
R-learning	2.05	0.02	1.25	0.97	0.39	0.87	75%	-0.1	3.5	0.06	0.06
				400	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	1.34	0.03	1.10	0.65	0.26	0.87	8%	0.0	3.1	0.37	0.30
MOM IPW	1.22	0.02	0.94	0.79	0.33	0.58	11%	0.0	3.2	0.34	0.10
MOM DR	1.04	0.02	0.87	0.78	0.32	0.45	0%	0.1	3.0	0.47	0.07
Causal Forest	0.97	0.02	0.84	0.70	0.30	0.52	8%	0.0	3.1	0.48	0.10
Causal Forest with local centering Lasso:	0.95	0.02	0.83	0.70	0.29	0.52	5%	0.0	3.1	0.48	0.10
Conditional mean regression	1.28	0.02	1.13	0.25	0.09	1.07	5%	0.0	3.0	0.47	0.61
MOM IPW	1.58	0.02 0.02	1.13	0.25 0.98	0.09 0.40	0.51	80%	-0.2	4.9	-0.01	0.01
MOM II W MOM DR	1.36	0.02 0.02	0.97	0.98 0.92	0.40 0.37	0.46	48%	-0.2	3.3	0.12	0.04
MCM MCM	1.63	0.02 0.02	1.09	1.02	0.36	0.48	69%	0.0	$\frac{3.3}{4.7}$	0.12 0.05	0.03
MCM with efficiency augmentation	1.03 1.39	0.02 0.02	0.98	0.93	0.38	0.45	53%	-0.1	3.3	0.03 0.21	0.03
R-learning	1.38	0.02 0.02	0.98	0.93	0.38	0.45	44%	-0.1	3.3	0.21 0.34	0.03
n-learning	1.58	0.02	0.98	0.93	0.38	0.45	44%	-0.1	ა.ა	0.54	0.03

Table D.23: Performance measures for GATE of ITE2 with random assignment and without random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	0 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	5.26	0.05	2.03	1.66	0.53	1.26	8%	0.0	3.0	0.66	0.19
MOM IPW	13.35	0.06	2.98	3.01	1.09	1.06	27%	0.0	3.1	0.56	0.04
MOM DR	10.96	0.06	2.59	2.74	0.99	0.92	39%	0.0	3.1	0.64	0.05
Causal Forest	8.07	0.07	2.22	2.21	0.77	1.15	59%	-0.1	3.1	0.66	0.11
Causal Forest with local centering	7.41	0.07	2.12	2.10	0.73	1.13	55%	-0.1	3.1	0.66	0.12
Lasso: Conditional mean regression	4.52	0.04	1.97	0.95	0.21	1.74	19%	0.0	3.1	0.65	0.33
MOM IPW	16.87	0.10	3.37	3.39	1.20	1.25	100%	-0.1	4.4	0.43	0.03
MOM DR	10.78	0.09	2.72	2.65	0.93	1.20	100%	-0.1	3.6	0.64	0.07
MCM	19.02	0.09	3.64	3.67	1.24	1.13	100%	-0.1	4.4	0.16	0.02
MCM with efficiency augmentation	11.17	0.09	2.79	2.71	0.96	1.18	100%	-0.1	3.4	0.60	0.06
R-learning	11.14	0.09	2.79	2.71	0.96	1.18	100%	-0.1	3.4	0.62	0.06
				400	0 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	1.70	0.03	1.01	0.73	0.13	0.88	17%	0.0	3.1	0.69	0.37
MOM IPW	8.08	0.09	2.14	2.36	0.84	0.67	11%	0.0	3.2	0.67	0.08
MOM DR	3.63	0.06	1.46	1.49	0.48	0.59	3%	0.1	3.0	0.69	0.19
Causal Forest	1.97	0.03	1.19	1.02	0.23	0.67	5%	0.1	3.0	0.70	0.30
Causal Forest with local centering	1.74	0.03	1.18	0.96	0.18	0.65	3%	0.1	3.0	0.70	0.32
Lasso:											
Conditional mean regression	1.45	0.02	1.17	0.44	0.02	1.05	11%	0.0	3.0	0.70	0.41
MOM IPW	8.95	0.15	2.47	2.47	0.88	0.86	47%	-0.1	3.3	0.66	0.08
MOM DR	3.49	0.05	1.51	1.48	0.49	0.69	25%	0.0	3.2	0.69	0.19
MCM	14.26	0.14	3.18	3.23	1.10	0.71	83%	0.0	3.8	0.58	0.02
MCM with efficiency augmentation	3.73	0.06	1.56	1.54	0.52	0.68	25%	0.0	3.2	0.69	0.18
R-learning	3.74	0.06	1.56	1.54	0.53	0.69	22%	0.0	3.2	0.69	0.18

D.2.4 GATEs from ITE with random assignment and random noise

The relative performances for the GATE estimators for ITEs with noise but randomized treatment assignment are very close to their selective equivalents. The only exception is ITE2 with 4,000 observations where the two Causal Forest versions perform best instead of MCM in the selective case. The latter is in the randomized setting the worst estimator. This emphasizes again that the averaging of noisy estimators is not always successful.

Table D.24: Performance measures for GATE of ITE0 with random assignment and random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	1.43	0.02	1.20	0.08	0.00	1.19	20%	0.0	3.1	0.00	1.02
MOM IPW	1.15	0.02	1.06	0.08	-0.01	1.06	30%	0.0	3.1	0.00	0.63
MOM DR	0.76	0.02	0.87	0.08	-0.01	0.86	8%	0.0	3.1	0.01	0.26
Causal Forest	0.95	0.02	0.97	0.08	-0.01	0.97	13%	0.0	3.1	0.00	0.37
Causal Forest with local centering $Lasso:$	0.89	0.02	0.94	0.08	0.00	0.94	6%	0.0	3.0	0.01	0.37
Conditional mean regression	2.87	0.03	1.63	0.08	0.00	1.67	20%	0.0	3.1	0.00	2.86
MOM IPW	1.09	0.03	1.01	0.08	-0.01	1.03	100%	0.1	4.4	0.01	0.54
MOM DR	0.79	0.02	0.88	0.09	0.00	0.88	63%	0.0	3.6	0.04	0.21
MCM	1.06	0.03	0.97	0.09	-0.04	1.02	100%	0.0	4.2	0.01	0.50
MCM with efficiency augmentation	0.75	0.02	0.86	0.09	0.00	0.86	61%	0.0	3.5	0.02	0.17
R-learning	0.74	0.02	0.86	0.09	0.00	0.85	58%	0.0	3.5	0.02	0.17
				400	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	0.77	0.02	0.84	0.09	0.01	0.86	11%	0.0	3.1	0.00	0.80
MOM IPW	0.35	0.01	0.59	0.08	0.01	0.58	11%	0.0	3.2	0.01	0.25
MOM DR	0.21	0.01	0.45	0.08	0.00	0.44	0%	0.1	3.0	0.01	0.10
Causal Forest	0.27	0.01	0.51	0.08	0.01	0.50	3%	0.1	3.0	0.01	0.15
Causal Forest with local centering	0.25	0.01	0.50	0.08	0.01	0.49	5%	0.1	3.0	0.01	0.15
Lasso:											
Conditional mean regression	1.12	0.02	1.03	0.09	0.01	1.04	9%	0.0	3.1	0.01	1.24
MOM IPW	0.24	0.01	0.47	0.08	0.01	0.48	63%	0.1	4.4	0.03	0.11
MOM DR	0.18	0.01	0.42	0.09	0.00	0.41	6%	0.1	3.0	0.03	0.04
MCM	0.23	0.01	0.46	0.09	-0.02	0.47	59%	0.1	4.4	0.00	0.10
MCM with efficiency augmentation	0.17	0.01	0.41	0.09	0.00	0.40	22%	0.1	3.1	0.03	0.03
R-learning	0.17	0.01	0.41	0.08	0.00	0.40	13%	0.1	3.0	0.06	0.03

Table D.25: Performance measures for GATE of ITE1 with random assignment and random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	1.55	0.02	1.23	0.32	0.12	1.18	17%	0.1	3.1	0.02	0.26
MOM IPW	1.27	0.02	1.09	0.32	0.11	1.06	23%	0.0	3.1	0.02	0.16
MOM DR	0.89	0.02	0.91	0.32	0.11	0.86	5%	0.0	3.0	0.03	0.07
Causal Forest	1.07	0.02	1.01	0.31	0.11	0.96	6%	0.0	3.1	0.03	0.09
Causal Forest with local centering	1.02	0.02	0.99	0.31	0.11	0.94	6%	0.0	3.0	0.03	0.10
Lasso:											
Conditional mean regression	2.93	0.03	1.65	0.26	0.09	1.66	19%	0.0	3.1	0.02	0.72
MOM IPW	1.25	0.03	1.07	0.35	0.12	1.03	100%	0.0	4.4	0.01	0.14
MOM DR	0.98	0.02	0.95	0.36	0.13	0.88	64%	0.0	3.6	-0.06	0.05
MCM	1.24	0.03	1.04	0.37	0.07	1.01	100%	0.0	4.4	0.00	0.13
MCM with efficiency augmentation	0.94	0.02	0.93	0.37	0.13	0.86	59%	0.0	3.5	-0.04	0.05
R-learning	0.93	0.02	0.92	0.36	0.13	0.86	64%	0.0	3.5	-0.09	0.04
				400	00 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	0.89	0.02	0.92	0.29	0.12	0.86	11%	0.0	3.1	0.03	0.21
MOM IPW	0.48	0.01	0.65	0.31	0.12	0.58	13%	0.0	3.1	0.03	0.07
MOM DR	0.34	0.01	0.52	0.31	0.11	0.45	0%	0.1	3.0	0.05	0.03
Causal Forest	0.39	0.01	0.58	0.29	0.11	0.51	3%	0.1	3.0	0.06	0.04
Causal Forest with local centering	0.38	0.01	0.57	0.29	0.11	0.50	3%	0.1	3.0	0.06	0.04
Lasso:											
Conditional mean regression	1.19	0.02	1.07	0.19	0.04	1.05	6%	0.0	3.0	0.06	0.34
MOM IPW	0.43	0.01	0.60	0.37	0.14	0.48	70%	0.1	4.3	-0.08	0.03
MOM DR	0.36	0.01	0.52	0.36	0.14	0.41	14%	0.0	3.0	-0.10	0.01
MCM	0.42	0.01	0.60	0.37	0.09	0.47	59%	0.1	4.3	0.00	0.03
MCM with efficiency augmentation	0.36	0.01	0.52	0.36	0.13	0.40	14%	0.1	3.1	-0.10	0.01
R-learning	0.36	0.01	0.51	0.37	0.14	0.40	17%	0.1	3.1	-0.04	0.01

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Table D.26: Performance measures for GATE of ITE2 with random assignment and random noise

	\overline{MSE}	$SE(\overline{MSE})$	Median MSE	$ \overline{Bias} $	\overline{Bias}	\overline{SD}	JB	Skew.	Kurt.	Corr.	Var. ratio
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)
				100	0 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	2.89	0.03	1.47	1.04	0.41	1.16	14%	0.0	3.1	0.06	0.02
MOM IPW	2.82	0.03	1.38	1.07	0.40	1.05	17%	0.0	3.1	0.05	0.01
MOM DR	2.44	0.02	1.24	1.07	0.39	0.86	6%	0.0	3.0	0.07	0.01
Causal Forest	2.51	0.03	1.29	1.03	0.38	0.96	11%	0.0	3.1	0.08	0.01
Causal Forest with local centering	2.46	0.03	1.28	1.02	0.39	0.94	8%	0.0	3.0	0.08	0.01
Lasso:											
Conditional mean regression	3.83	0.04	1.85	0.81	0.31	1.67	13%	0.0	3.0	0.08	0.07
MOM IPW	3.18	0.03	1.47	1.19	0.43	1.04	100%	-0.1	4.6	-0.07	0.01
MOM DR	2.85	0.03	1.38	1.17	0.43	0.90	77%	0.0	3.5	-0.01	0.01
MCM	3.15	0.03	1.51	1.22	0.27	1.01	100%	0.0	4.2	0.01	0.01
MCM with efficiency augmentation	2.83	0.03	1.37	1.18	0.44	0.88	72%	-0.1	3.5	-0.07	0.01
R-learning	2.81	0.03	1.37	1.17	0.43	0.88	77%	0.0	3.5	-0.06	0.01
				400	0 obse	rvatio	ns				
Random Forest:											
Conditional mean regression	1.86	0.03	1.13	0.89	0.36	0.84	9%	0.0	3.1	0.12	0.03
MOM IPW	1.87	0.02	1.00	1.01	0.39	0.58	8%	0.0	3.1	0.10	0.01
MOM DR	1.67	0.02	0.91	0.99	0.36	0.45	0%	0.1	3.0	0.13	0.01
Causal Forest	1.58	0.02	0.88	0.93	0.35	0.53	9%	0.0	3.1	0.14	0.01
Causal Forest with local centering	1.54	0.02	0.87	0.92	0.34	0.52	3%	0.0	3.0	0.14	0.01
Lasso:											
Conditional mean regression	1.59	0.03	1.19	0.55	0.18	1.04	5%	0.0	3.0	0.15	0.05
MOM IPW	2.33	0.02	1.16	1.18	0.44	0.51	84%	-0.3	4.8	-0.07	0.00
MOM DR	2.06	0.03	1.09	1.11	0.41	0.47	66%	0.0	3.4	0.00	0.00
MCM	2.38	0.02	1.22	1.23	0.29	0.48	67%	-0.1	4.7	-0.05	0.00
MCM with efficiency augmentation	2.09	0.02	1.09	1.12	0.41	0.46	72%	0.0	3.6	-0.01	0.00
R-learning	2.09	0.02	1.08	1.12	0.41	0.46	63%	0.0	3.6	-0.03	0.00

D.3 Results for ATE estimation

The Appendices D.3.1 to D.3.4 show the full results for ATE estimation in the 24 DGP-sample size combinations. Compared to IATEs and GATEs, the ATE performance measures require no averaging over several validation observations and provides the standard MSE, bias and SD for a point estimate. Also the summary of the fraction of observations with rejected JB test is not applicable in this case. Instead, we provide the p-value of the JB test to investigate whether the ATE estimators constructed as the average of the different IATEs are normally distributed.

The findings are similar to the findings for the GATE estimation. The differences in SD over all ATE estimators are minor making the bias the decisive component. This means that the estimators that account best for the selection bias perform best in the settings with selectivity. Like for the GATEs, we observe that averaging noisy IATE estimates can provide competitive ATE estimators. In particular, MCM shows consistently small bias for the ATE. As the bias drops close to zero in the randomized settings, there remains hardly any difference between the estimators.

D.3.1 ATEs from ITE with selection and without random noise

Table D.27: Performance measures for ATE of ITE0 with selection and without random noise (baseline)

	MSE	Bias	SD	Skew.	Kurt.	p-value JB
	(1)	(2)	(3)	(4)	(5)	(6)
			1000	observ	ations	
Random Forest:						
Conditional mean regression	0.97	0.60	0.78	-0.1	3.0	0.31
MOM IPW	1.18	0.71	0.82	0.0	3.1	0.34
MOM DR	0.68	0.40	0.72	-0.1	3.1	0.21
Causal Forest	1.20	0.75	0.80	-0.1	3.1	0.25
Causal Forest with local centering	0.77	0.34	0.81	0.0	3.0	0.39
Lasso:						
Conditional mean regression	1.00	0.60	0.80	0.0	3.0	0.42
MOM IPW	1.10	0.61	0.86	0.0	3.2	0.07
MOM DR	38.75	0.60	6.20	43.4	1921.8	0.00
MCM	0.94	0.45	0.86	-0.1	3.1	0.08
MCM with efficiency augmentation	0.87	0.42	0.83	0.0	3.1	0.24
R-learning	0.88	0.45	0.82	0.0	3.1	0.21
			4000	observ	ations	
Random Forest:						
Conditional mean regression	0.43	0.53	0.38	0.1	2.7	0.14
MOM IPW	0.46	0.53	0.41	0.2	2.8	0.07
MOM DR	0.21	0.28	0.37	0.1	2.8	0.17
Causal Forest	0.66	0.70	0.40	0.1	2.6	0.06
Causal Forest with local centering	0.22	0.24	0.40	0.1	2.8	0.24
Lasso:						
Conditional mean regression	0.39	0.49	0.39	0.0	2.9	0.39
MOM IPW	0.37	0.45	0.41	0.2	3.0	0.13
MOM DR	0.26	0.31	0.41	0.1	2.9	0.36
MCM	0.29	0.34	0.42	0.2	2.9	0.17
MCM with efficiency augmentation	0.23	0.26	0.40	0.1	3.0	0.38
R-learning	0.24	0.28	0.41	0.1	3.0	0.29

Table D.28: Performance measures for ATE of ITE1 with selection and without random noise $\,$

	MSE	Bias	SD	Skew.	Kurt.	p-value JB
	(1)	(2)	(3)	(4)	(5)	(6)
			1000	observ	ations	
Random Forest:						
Conditional mean regression	1.94	1.16	0.77	-0.1	3.0	0.32
MOM IPW	2.01	1.16	0.81	-0.1	3.1	0.20
MOM DR	1.06	0.75	0.70	-0.1	3.1	0.14
Causal Forest	2.48	1.36	0.79	0.0	3.1	0.30
Causal Forest with local centering Lasso:	1.16	0.72	0.80	0.0	3.0	0.38
Conditional mean regression	1.67	1.01	0.80	0.0	3.0	0.40
MOM IPW	1.61	0.94	0.85	-0.1	3.2	0.02
MOM DR	35.55	0.87	5.90	43.3	1913.6	0.00
MCM	1.06	0.57	0.85	-0.1	3.2	0.04
MCM with efficiency augmentation	1.17	0.70	0.83	0.0	3.2	0.13
R-learning	1.30	0.79	0.82	0.0	3.2	0.12
			4000	observ	ations	
Random Forest:						
Conditional mean regression	1.23	1.04	0.38	0.1	2.7	0.10
MOM IPW	0.83	0.82	0.41	0.1	2.8	0.15
MOM DR	0.38	0.50	0.36	0.1	2.9	0.23
Causal Forest	1.83	1.29	0.39	0.1	2.6	0.06
Causal Forest with local centering	0.46	0.55	0.40	0.0	2.8	0.27
Lasso:						
Conditional mean regression	0.68	0.73	0.40	0.0	2.8	0.35
MOM IPW	0.61	0.67	0.41	0.2	3.0	0.10
MOM DR	0.38	0.47	0.40	0.1	2.9	0.30
MCM	0.33	0.39	0.42	0.2	3.1	0.13
MCM with efficiency augmentation	0.34	0.43	0.40	0.1	3.0	0.29
R-learning	0.44	0.53	0.40	0.1	3.1	0.25

Table D.29: Performance measures for ATE of ITE2 with selection and without random noise ${\cal C}$

	MSE	Bias	SD	Skew.	Kurt.	p-value JB			
	(1)	(2)	(3)	(4)	(5)	(6)			
	1000 observations								
Random Forest:									
Conditional mean regression	5.62	2.24	0.76	0.0	3.1	0.31			
MOM IPW	5.70	2.25	0.79	-0.1	3.0	0.33			
MOM DR	3.17	1.64	0.68	-0.1	3.0	0.18			
Causal Forest	8.72	2.85	0.77	0.0	3.1	0.35			
Causal Forest with local centering	3.13	1.59	0.77	0.0	3.0	0.42			
Lasso:									
Conditional mean regression	2.75	1.46	0.79	0.0	3.0	0.41			
MOM IPW	3.85	1.77	0.85	-0.2	3.3	0.00			
MOM DR	36.56	1.32	5.90	43.3	1912.9	0.00			
MCM	1.68	0.97	0.86	-0.1	3.1	0.17			
MCM with efficiency augmentation	2.35	1.30	0.80	-0.1	3.3	0.02			
R-learning	3.00	1.53	0.80	0.0	3.1	0.29			
			4000	observa	ations				
Random Forest:									
Conditional mean regression	2.79	1.62	0.39	0.1	2.8	0.13			
MOM IPW	2.31	1.47	0.40	0.1	2.8	0.13			
MOM DR	1.07	0.97	0.35	0.1	2.8	0.28			
Causal Forest	6.24	2.47	0.40	0.1	2.6	0.05			
Causal Forest with local centering	1.08	0.97	0.38	0.1	2.8	0.14			
Lasso:									
Conditional mean regression	0.84	0.84	0.37	0.1	2.8	0.22			
MOM IPW	1.56	1.18	0.41	0.2	3.0	0.04			
MOM DR	0.52	0.62	0.38	0.1	2.9	0.19			
MCM	0.56	0.62	0.42	0.2	3.0	0.16			
MCM with efficiency augmentation	0.65	0.71	0.38	0.2	3.0	0.10			
R-learning	0.90	0.87	0.39	0.2	3.1	0.08			

D.3.2 ATEs from ITE with selection and random noise

Table D.30: Performance measures for ATE of ITE0 with selection and random noise

	MSE	Bias	SD	Skew.	Kurt.	p-value JB			
	(1)	(2)	(3)	(4)	(5)	(6)			
	1000 observations								
Random Forest:									
Conditional mean regression	0.96	0.60	0.78	-0.1	3.1	0.27			
MOM IPW	1.17	0.71	0.82	-0.1	3.1	0.22			
MOM DR	0.66	0.39	0.71	-0.1	3.1	0.22			
Causal Forest	1.18	0.74	0.79	0.0	3.1	0.35			
Causal Forest with local centering Lasso:	0.76	0.33	0.80	0.0	3.1	0.29			
Conditional mean regression	0.98	0.60	0.79	0.0	3.0	0.39			
MOM IPW	1.08	0.60	0.85	-0.1	3.2	0.05			
MOM DR	60.69	0.63	7.77	43.9	1949.8	0.00			
MCM	0.91	0.42	0.86	-0.1	3.2	0.05			
MCM with efficiency augmentation	0.87	0.42	0.83	0.0	3.1	0.23			
R-learning	0.88	0.45	0.82	0.0	3.2	0.20			
			4000	observ	ations				
Random Forest:									
Conditional mean regression	0.43	0.53	0.38	0.1	2.7	0.12			
MOM IPW	0.45	0.53	0.42	0.2	2.7	0.07			
MOM DR	0.21	0.27	0.37	0.1	2.8	0.16			
Causal Forest	0.64	0.70	0.39	0.1	2.7	0.09			
Causal Forest with local centering	0.21	0.23	0.40	0.1	2.8	0.19			
Lasso:									
Conditional mean regression	0.39	0.49	0.39	0.1	2.8	0.35			
MOM IPW	0.37	0.45	0.41	0.2	2.9	0.15			
MOM DR	0.26	0.31	0.40	0.1	2.9	0.27			
MCM	0.27	0.31	0.42	0.2	3.0	0.16			
MCM with efficiency augmentation	0.23	0.26	0.40	0.1	3.0	0.30			
R-learning	0.24	0.28	0.41	0.1	3.0	0.22			

Table D.31: Performance measures for ATE of ITE1 with selection and random noise (baseline)

	MSE	Bias	SD	Skew.	Kurt.	p-value JB			
	(1)	(2)	(3)	(4)	(5)	(6)			
	1000 observations								
Random Forest:									
Conditional mean regression	1.24	0.80	0.77	-0.1	3.1	0.24			
MOM IPW	1.42	0.87	0.81	0.0	3.1	0.31			
MOM DR	0.77	0.52	0.71	0.0	3.1	0.22			
Causal Forest	1.55	0.96	0.79	-0.1	3.1	0.24			
Causal Forest with local centering	0.86	0.48	0.80	0.0	3.1	0.31			
Lasso:									
Conditional mean regression	1.20	0.76	0.79	0.0	3.0	0.44			
MOM IPW	1.25	0.73	0.85	-0.1	3.3	0.00			
MOM DR	38.89	0.70	6.20	43.4	1922.0	0.00			
MCM	0.94	0.46	0.85	-0.1	3.2	0.05			
MCM with efficiency augmentation	0.95	0.52	0.83	0.0	3.2	0.15			
R-learning	1.00	0.58	0.82	0.0	3.2	0.17			
			4000	observ	ations				
Random Forest:									
Conditional mean regression	0.67	0.73	0.38	0.1	2.7	0.15			
MOM IPW	0.57	0.63	0.42	0.1	2.8	0.14			
MOM DR	0.26	0.35	0.37	0.1	2.9	0.27			
Causal Forest	0.99	0.91	0.40	0.1	2.7	0.11			
Causal Forest with local centering	0.29	0.35	0.40	0.1	2.8	0.26			
Lasso:									
Conditional mean regression	0.52	0.60	0.39	0.0	2.9	0.42			
MOM IPW	0.44	0.53	0.41	0.2	3.0	0.12			
MOM DR	0.30	0.37	0.40	0.1	3.0	0.34			
MCM	0.28	0.32	0.42	0.2	3.1	0.16			
MCM with efficiency augmentation	0.27	0.33	0.40	0.1	3.0	0.36			
R-learning	0.31	0.38	0.40	0.1	3.0	0.32			

Table D.32: Performance measures for ATE of ITE2 with selection and random noise (baseline)

	MSE	Bias	SD	Skew.	Kurt.	p-value JB		
	(1)	(2)	(3)	(4)	(5)	(6)		
	1000 observations							
Random Forest:								
Conditional mean regression	2.13	1.24	0.77	-0.1	3.1	0.24		
MOM IPW	2.15	1.23	0.81	-0.1	3.0	0.31		
MOM DR	1.17	0.82	0.71	-0.1	3.1	0.05		
Causal Forest	2.64	1.43	0.78	-0.1	3.1	0.23		
Causal Forest with local centering	1.29	0.81	0.80	-0.1	3.0	0.23		
Lasso:								
Conditional mean regression	1.88	1.12	0.80	-0.1	3.0	0.31		
MOM IPW	1.75	1.02	0.84	-0.1	3.2	0.01		
MOM DR	23.44	0.92	4.75	42.5	1867.4	0.00		
MCM	0.99	0.52	0.85	-0.1	3.2	0.02		
MCM with efficiency augmentation	1.29	0.79	0.82	-0.1	3.2	0.07		
R-learning	1.46	0.89	0.81	-0.1	3.1	0.12		
			4000	observa	ations			
Random Forest:								
Conditional mean regression	1.41	1.12	0.38	0.0	2.7	0.16		
MOM IPW	0.94	0.88	0.41	0.0	2.8	0.31		
MOM DR	0.44	0.56	0.36	0.0	2.9	0.39		
Causal Forest	1.99	1.36	0.39	0.1	2.7	0.18		
Causal Forest with local centering	0.55	0.63	0.39	0.0	2.9	0.42		
Lasso:								
Conditional mean regression	0.83	0.82	0.39	0.0	2.8	0.28		
MOM IPW	0.71	0.74	0.41	0.1	3.0	0.26		
MOM DR	0.46	0.54	0.41	0.1	3.1	0.37		
MCM	0.29	0.34	0.42	0.3	3.0	0.03		
MCM with efficiency augmentation	0.42	0.51	0.40	0.1	2.9	0.29		
R-learning	0.56	0.63	0.40	0.1	2.9	0.27		

D.3.3 ATEs from ITE with random assignment and without random noise

Table D.33: Performance measures for ATE of ITE0 with random assignment and without random noise

	MSE	Bias	SD	Skew.	Kurt.	p-value JB			
	(1)	(2)	(3)	(4)	(5)	(6)			
	1000 observations								
Random Forest:									
Conditional mean regression	0.62	0.01	0.79	0.0	3.1	0.38			
MOM IPW	0.66	0.01	0.81	0.0	3.1	0.23			
MOM DR	0.55	0.01	0.74	0.0	3.0	0.36			
Causal Forest	0.66	0.01	0.81	0.0	3.1	0.23			
Causal Forest with local centering	0.60	0.02	0.78	0.0	3.0	0.47			
Lasso:									
Conditional mean regression	0.62	0.01	0.79	0.0	3.1	0.40			
MOM IPW	0.67	0.01	0.82	0.0	3.1	0.29			
MOM DR	0.63	0.01	0.79	0.0	3.0	0.48			
MCM	0.67	0.01	0.82	0.0	3.1	0.26			
MCM with efficiency augmentation	0.61	0.02	0.78	0.0	3.0	0.44			
R-learning	0.61	0.02	0.78	0.0	3.0	0.43			
			4000	observa	ations				
Random Forest:									
Conditional mean regression	0.14	0.02	0.37	0.1	2.9	0.30			
MOM IPW	0.14	0.03	0.38	0.1	2.9	0.18			
MOM DR	0.12	0.02	0.35	0.1	2.9	0.29			
Causal Forest	0.15	0.03	0.38	0.1	2.9	0.33			
Causal Forest with local centering	0.13	0.02	0.37	0.1	2.9	0.26			
Lasso:									
Conditional mean regression	0.14	0.02	0.37	0.1	2.9	0.39			
MOM IPW	0.15	0.03	0.38	0.2	3.0	0.18			
MOM DR	0.14	0.02	0.37	0.1	2.8	0.29			
MCM	0.15	0.03	0.38	0.2	2.9	0.17			
MCM with efficiency augmentation	0.14	0.02	0.37	0.1	2.8	0.27			
R-learning	0.14	0.02	0.37	0.1	2.8	0.28			

Table D.34: Performance measures for ATE of ITE1 with random assignment and without random noise

	MSE	Bias	SD	Skew.	Kurt.	p-value JB			
	(1)	(2)	(3)	(4)	(5)	(6)			
	1000 observations								
Random Forest:									
Conditional mean regression	0.62	0.05	0.78	0.0	3.1	0.30			
MOM IPW	0.65	0.05	0.81	0.0	3.1	0.17			
MOM DR	0.54	0.05	0.73	0.0	3.1	0.35			
Causal Forest	0.65	0.05	0.81	0.0	3.1	0.19			
Causal Forest with local centering	0.59	0.06	0.77	0.0	3.0	0.46			
Lasso:									
Conditional mean regression	0.61	0.02	0.78	0.0	3.0	0.42			
MOM IPW	0.66	0.03	0.81	0.0	3.1	0.26			
MOM DR	0.62	0.03	0.79	0.0	3.0	0.47			
MCM	0.66	-0.02	0.81	0.1	3.1	0.23			
MCM with efficiency augmentation	0.61	0.04	0.78	0.0	3.1	0.40			
R-learning	0.60	0.04	0.77	0.0	3.1	0.38			
			4000	observa	tions				
Random Forest:									
Conditional mean regression	0.13	0.06	0.36	0.1	3.0	0.40			
MOM IPW	0.15	0.07	0.38	0.1	3.0	0.27			
MOM DR	0.12	0.05	0.34	0.1	2.9	0.29			
Causal Forest	0.14	0.06	0.37	0.1	2.9	0.34			
Causal Forest with local centering	0.13	0.06	0.36	0.1	3.0	0.36			
Lasso:									
Conditional mean regression	0.13	0.03	0.36	0.1	3.0	0.41			
MOM IPW	0.15	0.05	0.38	0.1	3.0	0.25			
MOM DR	0.13	0.04	0.36	0.1	2.8	0.33			
MCM	0.14	-0.01	0.38	0.1	3.0	0.21			
MCM with efficiency augmentation	0.14	0.04	0.37	0.1	2.9	0.31			
R-learning	0.14	0.04	0.37	0.1	2.9	0.30			

Table D.35: Performance measures for ATE of ITE2 with random assignment and without random noise

	MSE	Bias	SD	Skew.	Kurt.	p-value JB		
	(1)	(2)	(3)	(4)	(5)	(6)		
	1000 observations							
Random Forest:								
Conditional mean regression	0.57	0.04	0.76	0.0	3.1	0.28		
MOM IPW	0.66	0.09	0.81	0.0	3.2	0.12		
MOM DR	0.52	0.08	0.72	0.0	3.1	0.25		
Causal Forest	0.63	0.06	0.79	0.1	3.1	0.20		
Causal Forest with local centering	0.56	0.06	0.74	0.0	3.1	0.27		
Lasso:								
Conditional mean regression	0.55	-0.02	0.74	0.0	3.0	0.46		
MOM IPW	0.66	0.03	0.81	0.1	3.1	0.07		
MOM DR	0.59	0.02	0.77	0.0	3.1	0.34		
MCM	0.67	-0.03	0.82	0.1	3.0	0.17		
MCM with efficiency augmentation	0.58	0.02	0.76	0.1	3.1	0.24		
R-learning	0.58	0.03	0.76	0.1	3.1	0.18		
			4000	observa	tions			
Random Forest:								
Conditional mean regression	0.12	0.01	0.35	0.1	3.0	0.37		
MOM IPW	0.14	0.08	0.37	0.1	3.0	0.28		
MOM DR	0.11	0.03	0.33	0.1	3.0	0.39		
Causal Forest	0.13	0.01	0.36	0.1	2.9	0.39		
Causal Forest with local centering	0.12	0.00	0.35	0.1	3.1	0.37		
Lasso:								
Conditional mean regression	0.12	-0.02	0.35	0.1	3.1	0.38		
MOM IPW	0.14	0.02	0.38	0.1	3.0	0.23		
MOM DR	0.12	-0.01	0.35	0.0	3.0	0.49		
MCM	0.14	-0.03	0.38	0.1	3.0	0.24		
MCM with efficiency augmentation	0.13	0.00	0.36	0.1	3.0	0.43		
R-learning	0.13	0.00	0.35	0.1	3.0	0.43		

D.3.4 ATEs from ITE with random assignment and random noise

Table D.36: Performance measures for ATE of ITE0 with random assignment and random noise

	MSE	Bias	SD	Skew.	Kurt.	p-value JB			
	(1)	(2)	(3)	(4)	(5)	(6)			
	1000 observations								
Random Forest:									
Conditional mean regression	0.62	0.02	0.79	0.0	3.1	0.40			
MOM IPW	0.66	0.01	0.81	0.0	3.2	0.18			
MOM DR	0.54	0.01	0.74	0.0	3.0	0.41			
Causal Forest	0.65	0.01	0.81	0.0	3.2	0.18			
Causal Forest with local centering	0.59	0.01	0.77	0.0	3.0	0.47			
Lasso:									
Conditional mean regression	0.61	0.01	0.78	0.0	3.1	0.40			
MOM IPW	0.66	0.01	0.81	0.0	3.1	0.25			
MOM DR	0.62	0.02	0.79	0.0	3.1	0.44			
MCM	0.67	-0.02	0.82	0.0	3.1	0.23			
MCM with efficiency augmentation	0.61	0.02	0.78	0.0	3.0	0.45			
R-learning	0.60	0.02	0.78	0.0	3.1	0.42			
			4000	observa	tions				
Random Forest:									
Conditional mean regression	0.13	0.03	0.36	0.1	2.9	0.32			
MOM IPW	0.14	0.03	0.38	0.1	2.9	0.23			
MOM DR	0.12	0.02	0.34	0.1	2.9	0.25			
Causal Forest	0.14	0.02	0.38	0.1	2.9	0.27			
Causal Forest with local centering	0.13	0.02	0.36	0.1	2.9	0.29			
Lasso:									
Conditional mean regression	0.13	0.02	0.37	0.1	2.9	0.39			
MOM IPW	0.14	0.03	0.38	0.1	2.9	0.20			
MOM DR	0.13	0.02	0.36	0.1	2.8	0.26			
MCM	0.15	0.00	0.38	0.2	2.9	0.17			
MCM with efficiency augmentation	0.13	0.02	0.37	0.1	2.8	0.29			
R-learning	0.13	0.02	0.37	0.1	2.8	0.28			

Table D.37: Performance measures for ATE of ITE1 with random assignment and random noise

	MSE	Bias	SD	Skew.	Kurt.	p-value JB		
	(1)	(2)	(3)	(4)	(5)	(6)		
	1000 observations							
Random Forest:								
Conditional mean regression	0.61	0.04	0.78	0.0	3.1	0.25		
MOM IPW	0.65	0.03	0.81	0.0	3.1	0.26		
MOM DR	0.54	0.03	0.73	0.0	3.1	0.32		
Causal Forest	0.64	0.03	0.80	0.0	3.1	0.24		
Causal Forest with local centering	0.59	0.03	0.77	0.0	3.0	0.46		
Lasso:								
Conditional mean regression	0.60	0.03	0.78	0.0	3.1	0.39		
MOM IPW	0.65	0.02	0.81	0.0	3.1	0.27		
MOM DR	0.62	0.03	0.79	0.0	3.0	0.49		
MCM	0.66	-0.04	0.81	0.0	3.1	0.26		
MCM with efficiency augmentation	0.60	0.03	0.78	0.0	3.0	0.43		
R-learning	0.60	0.03	0.77	0.0	3.0	0.42		
			4000	observa	tions			
Random Forest:								
Conditional mean regression	0.14	0.05	0.37	0.1	2.9	0.20		
MOM IPW	0.15	0.04	0.38	0.2	2.8	0.12		
MOM DR	0.12	0.04	0.34	0.1	2.9	0.26		
Causal Forest	0.14	0.04	0.38	0.1	2.9	0.21		
Causal Forest with local centering	0.13	0.04	0.36	0.1	3.0	0.36		
Lasso:								
Conditional mean regression	0.13	0.03	0.37	0.1	3.0	0.30		
MOM IPW	0.15	0.04	0.38	0.2	2.9	0.15		
MOM DR	0.13	0.04	0.36	0.1	2.8	0.23		
MCM	0.15	-0.02	0.38	0.2	2.9	0.17		
MCM with efficiency augmentation	0.14	0.03	0.37	0.1	2.8	0.25		
R-learning	0.14	0.03	0.37	0.1	2.8	0.24		

Table D.38: Performance measures for ATE of ITE2 with random assignment and random noise $\frac{1}{2}$

	MSE	Bias	SD	Skew.	Kurt.	p-value JB			
	(1)	(2)	(3)	(4)	(5)	(6)			
	1000 observations								
Random Forest:									
Conditional mean regression	0.61	0.12	0.77	0.0	3.1	0.28			
MOM IPW	0.65	0.12	0.80	0.0	3.2	0.12			
MOM DR	0.54	0.10	0.73	0.0	3.0	0.34			
Causal Forest	0.64	0.11	0.79	0.0	3.1	0.17			
Causal Forest with local centering	0.60	0.12	0.76	0.0	3.1	0.28			
Lasso:									
Conditional mean regression	0.60	0.11	0.77	0.0	3.1	0.39			
MOM IPW	0.65	0.10	0.80	0.1	3.1	0.18			
MOM DR	0.62	0.11	0.78	0.0	3.1	0.40			
MCM	0.66	-0.08	0.81	0.0	3.1	0.26			
MCM with efficiency augmentation	0.60	0.11	0.77	0.0	3.0	0.36			
R-learning	0.60	0.11	0.76	0.0	3.1	0.29			
			4000	observa	tions				
Random Forest:									
Conditional mean regression	0.15	0.13	0.36	0.1	2.7	0.11			
MOM IPW	0.15	0.12	0.37	0.2	2.8	0.10			
MOM DR	0.12	0.10	0.34	0.1	2.8	0.22			
Causal Forest	0.15	0.12	0.37	0.1	2.9	0.27			
Causal Forest with local centering	0.14	0.11	0.36	0.1	2.7	0.15			
Lasso:									
Conditional mean regression	0.14	0.09	0.36	0.1	2.8	0.18			
MOM IPW	0.15	0.12	0.37	0.1	2.9	0.17			
MOM DR	0.14	0.11	0.36	0.1	2.7	0.14			
MCM	0.15	-0.06	0.38	0.1	3.0	0.32			
MCM with efficiency augmentation	0.14	0.11	0.36	0.1	2.6	0.10			
R-learning	0.14	0.11	0.36	0.1	2.6	0.09			

D.4 Computation time

This appendix shows the average computation times (in seconds) of the different estimation approaches. We computed all our results on a SWITCHengines cloud with 8 cores and 8GB RAM. It is difficult to compare the computation times between Random Forests and Lasso, because they depend strongly on the selection of the tuning parameters. The Lasso becomes slow when it selects many variables. The Causal Forests and MOM approaches differ in the way how they estimate the nuisance parameters.

Table D.39: Average computation time of one replication in seconds

	ITE0 w/o noise	ITE1 w/ noise	ITE2 w/ noise
	(1)	(2)	(3)
	10	000 observation	s
Random Forest:			
Infeasible	1.1	2.6	2.8
Conditional mean regression	4.0	4.1	4.0
MOM IPW	5.2	5.1	5.2
MOM DR	8.2	8.2	8.1
Causal Forest	3.9	3.9	3.9
Causal Forest with local centering	5.2	5.2	5.2
Lasso:			
Infeasible	_	26.8	29.5
Conditional mean regression	7.6	7.7	7.7
MOM IPW	12.4	12.3	12.3
MOM DR	17.9	17.9	17.9
MCM	11.3	11.3	11.3
MCM with efficiency augmentation	17.4	17.4	17.4
R-learning	17.4	17.4	17.4
	40	000 observation	ıs
Random Forest:			
Infeasible	3.2	8.6	9.7
Conditional mean regression	11.2	11.4	11.3
MOM IPW	17.0	17.0	17.0
MOM DR	32.4	33.1	32.8
Causal Forest	11.6	11.8	11.7
Causal Forest with local centering	18.3	18.3	18.3
Lasso:			
Infeasible	_	40.5	46.4
Conditional mean regression	24.2	24.1	24.2
MOM IPW	49.6	49.4	49.2
MOM DR	68.0	67.9	67.9
MCM	51.8	51.7	51.5
MCM with efficiency augmentation	67.4	67.2	67.2
R-learning	67.4	67.2	67.3