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Do fences make good neighbours?

Evidence from an insurgency in India

Heidi Kaila,¹ Saurabh Singhal,² and Divya Tuteja³

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Abstract: India has employed a variety of military, political, and economic measures to combat the long running insurgency in Kashmir with little evidence on what contributes to stability in the region. This paper uses a variety of tests to detect structural breaks in the time series for violence over the period 1998–2014. We identify a transition from a high violence regime to a low violence regime that coincides with (i) the fencing of the border with Pakistan (ii) the implementation of a large-scale development programme, and (iii) the phasing in of the Indian National Rural Employment Guarantee Scheme (NREGS). Our results highlight the complementary roles of development programmes and security in reducing violence.

Keywords: conflict, multiple structural breaks, Nonlinear Time Series Models, Jammu, Kashmir, India

JEL classification: C22, D74, F51

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1 Introduction

While the number of inter-state and civil wars have declined over time, close to 1.5 billion people remain affected by fragility, conflict and violence (World Bank, 2011). Undoubtedly, poverty and conflict go hand-in-hand and policymakers in developing countries continue to grapple with policies that could end the recurrent cycles of poverty and violence. In particular, researchers and policymakers have been demonstrating an increasing interest in examining the utility of development programs in the context of conflict reduction and the conditions necessary for their success. However, isolating the effect of such policies is complex, given that multiple policies maybe simultaneously implemented. In this paper, we utilize a variety of time series techniques to assess several policies implemented in the context of the ongoing insurgency in the Indian state of Jammu and Kashmir.

Beset with numerous insurgencies within its borders since independence, India has employed a combination of military, political and economic measures to combat them, yet there is little evidence on what factors may have contributed to stability. One such case is that of the ongoing conflict in the Indian state of Jammu and Kashmir. Both India and Pakistan claim territorial control over Kashmir with each currently controlling two-thirds and one-third of the area, respectively. The current insurgency started in 1989 with separatists, backed by Pakistan, contesting control of the Indian government (Chadha, 2005; Habibullah, 2008). A key security policy implemented to counter the insurgency was fencing of the Line of Control (LoC), the de facto border between Indian and Pakistan controlled parts of Kashmir, by the end of 2004. This was followed by the introduction of two large-scale development programs - the Indian National Rural Employment Guarantee Scheme (NREGS) and the Prime Minister's Reconstruction Plan for J&K (PMRP) in the state.

While understanding the linkages between conflict and socio-economic outcomes has long been considered important, there has been a recent shift in focus towards the effects of counterinsurgency policies in the last decade primarily due to the wars in Iraq and Afghanistan (Berman and Matanock, 2015). Although government forces are typically better equipped than insurgents, civilian support often plays a large part in successful operations. Cognizant

of this, the “hearts and minds” approach aims to win over the population by providing them public services, with the expectation that once their grievances are addressed, the attitude of the population towards the government will improve. The civilians are then less likely to help or join the insurgents and more likely to share information with the counterinsurgents. Berman et al. (2011b) find that improved service provision through the Commanders Emergency Reconstruction Program (CERP) in Iraq reduced violence, especially in the case of small-scale projects implemented in consultation with local leaders.

A second related counterinsurgency mechanism banks on the opportunity-cost mechanism approach, which posits that an improved economic environment, access to the market, labor market conditions, etc. increases the costs of participating in the insurgency and reduces the supply of insurgents. Miaari et al. (2014) examine how restrictions on the employment of Palestinians in Israel following the outbreak of the Second Intifada affected the involvement of Palestinians in the conflict. Exploiting spatial differences in the decline of employment opportunities for Palestinians, they find a 1% point decline in the employment rate to be associated with an increase of 0.11 Palestinian fatalities. Similarly, Iyengar et al. (2011) find that labor-intensive projects under the CERP reduced violence levels in Iraq.¹

On the other hand, it is also possible that such development programs could attract more violence through rent seeking or predatory behavior on the part of the insurgents. Insurgents may also try to sabotage developmental activities in an effort to undermine the government. For example, Crost et al. (2014) find that districts eligible for the KALAH-CIDSS development assistance program in the Philippines witnessed an increase in violence. The authors’ hypothesis is that this is due to insurgents strategically trying to sabotage projects. Similarly, Beath et al. (2013) find that even though the National Solidarity Program (NSP) in Afghanistan improved villagers’ perception of the government, its effect on security was temporary and dissipated over time.

Thus, a crucial determinant for the success of development programs in reducing violence may be a sufficiently low initial level of violence. This implies that security and development programs may complement each other, such that security policy must be effective first if

¹However, Berman et al.(2011a) find evidence against this mechanism.

initial violence levels are high. Indeed, Berman et al. (2013) argue that troop deployments and the resulting improvements in security played a crucial role in the success of the CERP program in Iraq. Emerging evidence from Afghanistan points to a similar relationship. Using geo-coded data, Sexton (2016) finds counter-insurgency aid reduced violence in areas that were already under the control of pro-government forces and Beath et al. (2013) find that the NSP program only had a positive effect on security in villages that had low initial levels of violence.² This paper further highlights the role played by improved security in the effectiveness of development programs in reducing violence. Using a variety of time series techniques, we find that fencing of the border between India and Pakistan reduced the level of violence in the Indian state of Jammu and Kashmir by restricting the supply of insurgents. Development programs implemented subsequently in the improved security environment further reduced violence, particularly in the form of civilian casualties.

In this paper, we use various endogenous structural break models to test if the fencing of the LoC and the implementation of the PMRP and NREGS lead to structural changes in the insurgency in Jammu and Kashmir. Even when the start date is known, policy interventions may affect outcomes gradually over time making it difficult to precisely *ex-ante* identify a break date in the outcome variable. We explore a smooth break in the time series by employing Logistic Smooth Transition Regression (LSTR) with time as the threshold variable, and find that the level of violence depicts a nonlinear break centered around the beginning of 2005, corresponding to the fencing of the LoC. Estimates from the Bai and Perron (1998, 2003) methodology to detect multiple unknown structural breaks further underscore this result. Our results are robust to different model specifications and transformations of the data. We find a significant structural change marked by a decline in the insurgency (particularly the number of terrorists killed) in 2005. This is followed by a decline in the number of civilian and security forces casualties in 2006-07, indicating that the improved security coupled with the introduction of two large-scale development programs, the PMRP and NREGS, further helped reduce violence in the state. This pattern in the

²In a recent systematic review of the existing literature, Zürcher (2017) finds that aid reduces violence conditional on there being a secure environment for the implementation of aid projects.

timing of breaks is indicative of the causal factors that may have been at play during the period of declining violence in the state.

This paper contributes to the literature examining conflict in India. In the Indian context, Muralidharan et al. (2017) find that NREGS increased earnings of low-income households, primarily via increase in private sector wages. However, the evidence on the effect of NREGS on the long-running Maoist violence in central and eastern India is mixed. While Khanna and Zimmermann (2017) find that it increased violence in the short-run, Dasgupta et al. (2017) and Fetzer (2014) find that it reduced violence levels most likely due to the significant rural poverty reductions associated with the program.³ Our results indicate that development programs are related to reduction in violence, albeit in the presence of improved security.

Finally, this paper also contributes to the use of nonlinear time series methods in the study of conflict. To our knowledge, the existing studies on conflict which are based on time series analysis only employ methods that detect sharp breaks, mainly using the Bai and Perron procedure (Bai and Perron, 1998; 2003).⁴ For example, Amara (2012) utilizes a combination of endogenous sharp break models along with exogenous structural break tests of Chow (1960) and Quandt-Andrews (Andrews, 1993) to study the relationship between the U.S. military ‘surge’ and economic and security stability in Iraq. Similarly, using endogenous (but sharp) structural breaks, Enders and Sandler (2005) study incidents of transnational terrorism with a focus on the changes that may have been triggered by 9/11.⁵ We focus, instead, on endogenous smooth breaks inferred using the LSTR framework in addition to sharp breaks given by the Bai and Perron procedure. This is because ongoing conflicts are more likely to adjust gradually to a long-run equilibrium than exhibit sharp movements. The presence of non-linearities in our data allows us to gain a deeper understanding of the evolution of smooth breaks and violence in the state. It is noteworthy that both the smooth and sharp breaks are detected without *a priori* assuming the dates when they take place.

³In related work, Singhal and Nilakantan (2016) assess the effectiveness of a security policy implemented to combat the Maoist insurgency.

⁴An exception is Enders et al. (2016) that finds a nonlinear relationship between income level and terrorism.

⁵Other papers which have employed sharp breaks in the context of conflict studies include Oosterlinck (2003), Waldenström and Frey (2008), and Chaney (2008).

Our methodology, hence, lets “the data speak for itself”.

The rest of the paper is organized as follows. In Section 2 we provide an overview of the ongoing conflict in the Indian state of Jammu and Kashmir. Section 3 outlines the empirical strategy and the data used in the study. The results are presented in Section 4 and Section 5 concludes.

2 The context: Insurgency in Jammu and Kashmir

The low intensity conflict that started in 1989 is rooted in the dispute between India and Pakistan over the territory of Kashmir, ongoing since the partition of the Indian subcontinent in 1947. Currently, India and Pakistan control two-thirds and one-third of the original state of Jammu and Kashmir, respectively. The dispute has led to two open wars, in 1947 and 1965, and brought the two countries close to war on a number of other occasions. A map of the region is provided in Figure 1.

The current armed insurgency started in 1989 in the Kashmir Valley, spreading over time to other parts of the state. A variety of factors contributed to the rise and spread of the violent insurgency, including widespread discontent with elections in the state and active support from Pakistan in the form of arms and training (Business Standard, 2015).⁶ The Indian army was summoned to quell the insurgency and it continues to run the counter-insurgency operations in the state in conjunction with central and state police forces. During the period from 1998 to 2014, the insurgency resulted in over 25,000 deaths (Source: SATP). However, the number of casualties has reduced drastically since 2005. The average number of total people killed during the periods 1991-1995, 1996-2000, 2001-2005, 2006-2010 and 2011-2015 were 2313, 2672, 2724, 637 and 170 respectively (Source: SATP). In addition to the considerable loss of life, recent research also finds that children born during the conflict are smaller and complete lesser years of schooling (Parlow, 2011 and 2012).

Table 1 lists the important incidents related to the insurgency in Jammu and Kashmir.

⁶ The end of the Soviet occupation in neighboring Afghanistan in 1989 also resulted in the availability of excess arms and experienced fighters. For a more detailed discussion of the conflict in Jammu and Kashmir, see Chadha (2005).

Figure 2 displays the timeline of the events along with the total monthly casualties during the insurgency. One of the important incidents involved the Kargil war and its aftermath. In early spring of 1999, armed intruders were discovered to have taken over strategic positions on the Indian side of the LoC. The Indian army was mobilized and moved to the border areas to repel the intruders. The war ended shortly in July when the Indian army successfully repulsed the intruders and chose not to widen to conflict with Pakistan. The quick movement of the army to the border regions disrupted their regular counterinsurgency operations in the interior regions. As a result, the interior areas vacated by the army were occupied by insurgents. Following the end of the Kargil war, the army had to (re)contest for control of the interior regions leading to an increase in violence that only came down by 2003, when India and Pakistan restored diplomatic ties and agreed to a cease-fire along the LoC (BBC News, 2003).

The Indian government has used a blend of military, political and economic policies to combat the insurgency. On the political front, the government has engaged with Pakistan at various points of time without any apparent success (except for the 2003 cease-fire agreement that has held despite a few violations). For example, a ceasefire negotiated with the primary militant groups from November 2000 to May 2001 collapsed without making much headway. The most significant political change in the 2000s has been the successful implementation of state elections. The state was under the President's Rule (i.e., the central government) for most of the 1990s and the dominant regional political party Jammu Kashmir National Conference (JKNC) was viewed as corrupt. The 2002 elections were a watershed in the electoral history of the state, allowing the popular Jammu Kashmir Peoples Democratic Party (PDP) to take charge (in coalition with the Indian National Congress, INC).

One of the key security measures undertaken by the government has been fencing the border with Pakistan. In terms of economic interventions, the notable intervention in the state during this time period was the introduction of the Prime Minister's Reconstruction Plan for J&K (PMRP) in 2005. Following this, the Indian National Rural Employment Guarantee Scheme (NREGS) was rolled out in the state over the period 2006-08. These

policies are discussed in greater detail below.

2.1 Fencing of the Line of Control

The erstwhile princely state of Jammu and Kashmir is delineated into the Indian and Pakistan controlled parts by the “Line of Control (LoC)”.^{7,8} India has fenced its border with Pakistan, both international and the LoC (Waldman, 2004). Fencing of the LoC, around 550 kilometers of the 740 kilometers allowing for breaks in the terrain, was completed by September 2004 and Indian security forces estimate that it has been particularly successful in reducing the infiltration of militants from Pakistan (Times of India, 2004; The Indian Express, 2014). Figure A4 in the Appendix 4 displays a photograph of the fence.

2.2 Prime Minister’s Reconstruction Plan for J&K (PMRP)

The PMRP was announced by Prime Minister Dr. Manmohan Singh in November 2004. The objective of the plan was the long-term development of the state through the creation of infrastructure, provision of basic services, and creation of jobs. The infrastructure projects included within the ambit of the plan are expansion of the road network, power generation projects, rural electrification, construction of health centers and Anganwadis, and the construction of colleges.^{9,10} Support was provided for the tourism industry through modernization of airports, conservation programs for various lakes, construction of tourist villages and training support for those in the tourism and hospitality industry. Income and employment generation in the agricultural sector was supported through various programs in the

⁷This was originally called the Cease-fire line following the first war in 1947-48. It was re-designated as the “Line of Control” after the Shimla Agreement in 1972. While the Line of Control is not internationally recognized, it is considered the de-facto border between India and Pakistan.

⁸A small section of the border between Indian and Pakistan controlled parts is part of the internationally recognized border.

⁹ Anganwadis are government funded child-care centers in India.

¹⁰ Some of the rural electrification objectives are included under the Rajiv Gandhi Grameen Vidyutikaran Yojana (RGGVY), a large India-wide program launched in 2005 with the objective of providing electricity access to hitherto un-electrified villages.

horticultural industry and construction of food storage units.¹¹

As of August 2015, 36 out of 67 sanctioned projects had been completed. While the central government had allocated approximately 240 billion Rupees (or 4 billion USD at the exchange rate of 1 USD = 60 Rupees) for the initial four-year period (2005-08), as of March 2014, over 780 billion Rupees (13 billion USD) have been disbursed under the scheme. The yearly expenditures of the reconstruction plan are displayed in Figure 3.

2.3 Indian National Rural Employment Guarantee Scheme (NREGS)

The NREGS guarantees 100 days of manual work at the minimum wage to all rural households. The objective is to protect the livelihood of rural households in times of dire need and is considered to be one of the largest safety net programs in the world. The National Rural Employment Guarantee Act (NREGA) was passed into law in August 2005 and the NREGS was phased in, in a non-random manner, between 2006 and 2008. More specifically, the NREGS was rolled out in three phases in India: in Phase 1, 200 districts received the scheme beginning February 2006; in Phase 2, an additional 130 districts were added to the program starting April 2007; and finally, in Phase 3 all the remaining districts were covered in April 2008. In Jammu and Kashmir, 3 districts were covered under Phase 1 namely Doda, Kupwara and Poonch. In Phase 2, Anantnag and Jammu were added to the scheme and, finally, the scheme was operational in all districts of the state by the start of Phase 3.¹²

3 Empirical Methodology and Data

This section deals with the empirical methodology and data utilized in the study. First, motivated by the nonlinear nature of our data, we use the logistic smooth transition regression (LSTR) methods to discern a nonlinear structural change in the data. Thereafter, we use

¹¹ The plan also provided funds for the rehabilitation of families affected by militancy in the state. Further details are available at <<http://pib.nic.in/newsite/erelcontent.aspx?relid=4947>> (accessed Jan 18, 2016).

¹² NREGA was later renamed Mahatma Gandhi National Rural Employment Guarantee Act. Further details regarding the policy are available at <www.nrega.nic.in/> (accessed Jan 18, 2016).

Bai and Perron (1998, 2003) technique to detect multiple sharp structural breaks in keeping with the existing literature that utilizes time-series methods to identify structural breaks in conflicts (Enders and Sandler 2005; Amara 2012).

3.1 Logistic Smooth Transition Regression (LSTR)

Economic variables are likely to depict a structural change with gradual shift over time. One of the possible reasons for such a behavior could be a slow reaction of economic agents to policy measures. Therefore, in our analysis, we focus on the existence of smooth breaks in the data. In order to incorporate nonlinear breaks, we consider a logistic smooth transition regression model (LSTR) with time as the threshold variable (Teräsvirta, 1994; 1998; Lin and Teräsvirta, 1994; Enders, 2015).¹³

We now introduce the nonlinear framework for time series analysis, specifically the LSTR model. A nonlinear autoregressive model is a generalization of the standard time series model with autoregressive coefficients and can be depicted as follows:

$$y_t = \eta' z_t + \theta' z_t H(\gamma, c, t) + u_t; \quad t = 1, \dots, T \quad (1)$$

where $z_t = (y_{t-1}, \dots, y_{t-p})$ is the vector of past realizations of the dependent variable y_t (and could also include exogenous regressors); $\eta = (\eta_0, \eta_1, \dots, \eta_m)'$ and $\theta = (\theta_0, \theta_1, \dots, \theta_m)'$ denote the parameter vectors and u_t is the error term. The smooth, bounded and continuous transition function is given by $H(\gamma, c, t)$ with transition variable time, i.e. t . The smoothness parameter is denoted by γ and c is the centrality parameter. In this model, the AR decay is dependent on the value of time t as time is the threshold variable. As a result, the intercept and AR coefficients smoothly vary across the regimes.

¹³If the series display a smooth structural break which leads to a gradual movement across the regimes, we need a nonlinear framework, such as a smooth transition regression, in order to capture the smoothed break in such phenomenon. One of the plausible assumptions that we could impose to gauge such a nonlinear break in the series is to consider a logistic break with time as the threshold variable within the STR framework. To choose the appropriate model for capturing the behavior of our data, first we run nonlinearity tests to see whether a model in the STR family would be appropriate. After that, we study whether an LSTR model is appropriate by running an auxiliary regression. The detailed results of model selection are presented in section 4.

Further, the LSTR model assumes that the transition function governing the movement across the states is a logistic function, therefore, the above model takes the following form:

$$y_t = \alpha_0 + \alpha_1 y_{t-1} + \cdots + \alpha_p y_{t-p} + \theta [\lambda_0 + \lambda_1 y_{t-1} + \cdots + \lambda_p y_{t-p}] + u_t \quad (2)$$

where $\theta = [1 + \exp(-\gamma(t - c))]^{-1}$ with intercepts α_0 and λ_0 , AR coefficients $\alpha_1, \dots, \alpha_p$ and $\lambda_1, \dots, \lambda_p$ and with optimal lag length p .

Finally,

$$y_t = \begin{cases} \alpha_0 + \alpha_1 y_{t-1} + \cdots + \alpha_p y_{t-p} + u_t, & \text{if } t < c \\ \alpha_0 + \alpha_1 y_{t-1} + \cdots + \alpha_p y_{t-p} + \theta [\lambda_0 + \lambda_1 y_{t-1} + \cdots + \lambda_p y_{t-p}] + u_t, & \text{if } t \geq c \end{cases} \quad (3)$$

since $H_{-\infty} = 0$ as $t \rightarrow -\infty$, $\theta = 0$ and $H_{\infty} = 1$ as $t \rightarrow \infty$, $\theta = 1$.

Over time, the value of θ goes from 0 to 1 and the AR coefficients gradually shift from the first state to the second state.

In the present study, we specify a LSTR model such that the constant varies across the states and the model simplifies to:¹⁴

$$y_t = \alpha_0 + \alpha_1 y_{t-1} + \cdots + \alpha_p y_{t-p} + \lambda_0 [1 + \exp(-\gamma(t - c))]^{-1} + u_t \quad (4)$$

However, we need to test for the existence of nonlinear breaks or the appropriateness of the LSTR model with time as the threshold variable using an auxiliary regression assuming $\theta = [1 + \exp(-\gamma(t - c))]^{-1}$. Lin and Teräsvirta (1994) suggest adoption of a Taylor series expansion of θ and so we estimate the following auxiliary regression equation (Enders, 2015):

$$y_t = a_0 + a_1 t + a_2 t^2 + a_3 t^3 + \sum_{i=1}^p b_i y_{t-i} + \varepsilon_t \quad (5)$$

where, in order to test for the existence of a LSTR break, we test the null hypothesis that $a_1 = a_2 = a_3 = 0$.

¹⁴The constant and trend specifications were also attempted but the trend variable was insignificant.

3.2 Bai and Perron (BP) Methodology

The detection of structural breaks using the dummy variable approach, which has been standard in the literature, is essentially based on the assumption that the breaks are sharp and the impact on the variable of interest at a break date is observed immediately at that point of time.

The general formulation of the model with m breaks and $m + 1$ regimes is:

$$y_t = x_t' \beta + z_t' \delta_j + u_t, \quad t = T_{j-1} + 1, \dots, T_j \quad (6)$$

for $j = 1, \dots, m + 1$. In the model y_t is the observed dependent variable at time t ; x_t ($r \times 1$) and z_t ($q \times 1$) are vectors of covariates and β and δ_t denote the vectors of coefficients and u_t is the error term. The indices of the break points T_1, \dots, T_m are treated as unknown. This is called a general partial structural change model, since β is not subject to shifts and is estimated using the full sample. A pure structural break model occurs thus when $r = 0$.

Various test statistics can be employed to determine the number of breaks given a maximum number of breaks m .¹⁵ The $F(l + 1 | l)$ test statistic tests the hypothesis of $l + 1$ against l structural breaks and the $\sup F(k; q)$ test statistic tests the hypothesis of zero breaks against k breaks with q break parameters (i.e., endogenous regressors). The UDmax statistic tests the null hypothesis of no structural break against an unknown number of breaks.

To determine the exact number of breaks one can use several information criteria: BIC (Bayesian Information Criterion), LWZ (Modified Schwarz Criterion) or a sequential procedure based on the $(l + 1 | l)$ test statistic. Bai and Perron (2003) point out that when breaks are present, BIC performs well, while LWZ performs better under the null, and might underestimate the number of breaks when the null is rejected. A drawback of the sequential procedure, according to Enders (2015), is that it may perform poorly if the series is highly persistent. Since our data display high persistence and the UD max and $\sup F(k; q)$ statistic firmly reject the null, we use the BIC procedure to determine the number of breaks.

Our main results are obtained by estimating a model where the vector z_t' includes a

¹⁵Bai and Perron (2003) provide examples with $m = 5$.

constant and a trend, and $p = 0$, so we do not have exogenous regressors.

$$y_t = \delta_{1j} + \delta_{2j}t + u_t, \quad t = T_{j-1} + 1, \dots, T_j \quad (7)$$

As a robustness check, in the online appendix we also show results with AR-terms as exogenous regressors.¹⁶ The model then becomes:

$$y_t = \delta_{1j} + \delta_{2j}t + \beta_i \sum_{i=1}^p y_{t-i} + u_t, \quad t = T_{j-1} + 1, \dots, T_j \quad (8)$$

where p is the number of lags.

3.3 Data

As indicators of violence we use the number of casualties and the number of incidents involving the use of explosives (landmines, grenades, IEDs, etc.). Both are available at the monthly level and are collected from the website of the South Asian Terrorism Portal (SATP), which bases these estimates on newspaper reports on terrorism related incidents in Jammu and Kashmir.¹⁷ While the data on the number of incidents involving explosives are available from January 2001 to December 2014, the data on the number of casualties are available for the period January 1998 to December 2014. We are able to disaggregate the data on casualties further into the number of civilians, security personnel, and terrorists killed, which allows us to examine if the results are driven by violence against a particular group.

Additionally, we can differentiate foreign terrorists killed from the aggregate number of terrorists killed by using police records for the period January 2003-December 2010.¹⁸ While the records are available for the period October 1998-April 2010, no foreign terrorist casualties are reported for the period April 2002-June 2003. Given the number of casualties reported before and after this period, we regard this data as missing and only consider data from July 2003-December 2010 for our analysis. The descriptive statistics are reported in

¹⁶The selection of the number of lags k is presented alongside the results.

¹⁷See <http://www.satp.org/> for details.

¹⁸These data were obtained from the J&K police website (<http://www.jkpolice.gov.in/index.htm>) in 2014.

Table 2 and the time series are plotted in Figure 4 (Panels 4a to 4f).

Before proceeding to the time series analysis, we check for the presence of unit roots in all the series in the following way. First, we select the optimal lag length for the unit root tests. In order to do that, we utilize the conventional lag length criterion viz. Akaike Information Criterion (AIC), Bayesian Information Criterion (BIC), Hannan-Quinn (HQ) and F-statistic along with the lag exclusion tests to choose the appropriate lag length for each of the series.¹⁹ Thereafter, we test for the existence of a unit root in the time series. In view of suspected structural breaks in our data, we conduct the Lee and Strazicich test for the presence of a unit root with structural breaks and the results are given in Table 3.²⁰ The null hypothesis of a unit root process in the presence of structural breaks is rejected at 1% level of significance in all the cases. Therefore, we conclude that the series are stationary in levels and employ these for the analysis of structural breaks.

4 Results

This section presents results of the time series analysis of the casualties and incidents involving explosives during the insurgency in Jammu and Kashmir. To begin with, we investigate the existence of nonlinear smooth breaks and sharp breaks to understand the evolution of the conflict. This is followed by a brief discussion on the robustness checks.

4.1 Nonlinear Breaks

As discussed in Section 3.1, most economic time series depict a smoothed break, which takes place over a period of time instead of a sudden change. This is the case for instance, when economic variables respond slowly to a change in economic policy. In such scenarios, a smooth transition model can be applied which captures transitions from one economic

¹⁹Detailed results of the lag selection and lag exclusion tests are available with the authors on request.

²⁰Lee and Strazicich (2003) propose a two-break minimum LM unit root test with breaks in the level and trend under the null hypothesis which they argue, conclusively implies trend stationarity under the alternative hypothesis.

regime to another that is typical of nonlinear time series data. In order to capture smooth breaks, we resort to smooth transition regression models with time as the threshold variable.

First, we test if the violence series are indeed nonlinear using the tests proposed by Tsay (1986) and Luukkonen et al. (1988). The null hypothesis of both the tests is linearity. The alternative hypothesis of Tsay's test is nonlinearity and that of Luukkonen et al. (1988)'s test is a smooth transition autoregressive (STAR) model. From Table 4, we observe that both tests indicate that the series are nonlinear. In particular, results of the Luukkonen et al. (1988) test suggest that a STAR model may be appropriate. Next, we need to select the optimal lag for the nonlinear models. Following Franses and van Dijk (2000), we choose optimal lag lengths based on the information criterion (AIC, BIC and HQ) and then estimate the corresponding nonlinear models using the suggested lag lengths. We estimate the LSTR model with the optimal lag lengths and pare down lags in the model by dropping variables which are insignificant. We finally select one lag for the civilians killed, terrorist killed and total killed, two lags for security personnel killed, and one lag for incidents involving explosives.²¹

The estimates of the LSTR models for the series are given in Table 5. Table 5a presents the results for the auxiliary regression given in equation 5. The F-statistic corresponds to the null hypothesis of no smooth transition in the threshold variable, time. The null hypothesis is rejected in all the five cases at 1% level of significance, confirming the existence of a smooth transition regression process with time as the transition variable.

In Table 5b, we present the main results of the nonlinear model given in equation 4 which estimates the LSTR model with time as the threshold variable.²² It is notable that λ_0 is negative in all the cases which signifies that the mean of the process is decreasing as $\theta \rightarrow 1$. There seems to exist two states in all the time series on casualties: a high casualties state until 2003 and a low casualties state from 2007 onwards, and the period in between, from

²¹Due to the lack of data in the foreign terrorists killed series, we do not perform the LSTR analysis for this variable.

²²We also estimate the LSTR models with two smoothed breaks for all the series. The modified AIC statistic (Enders and Holt, 2012) is utilized to select the final model, $AIC = T \log \left(\sum_{t=1}^T \hat{u}_t^2 \right) + 2r$, where T is the number of time periods, r is the number of parameters estimated in the respective models and \hat{u}_t are the residuals of the model. The best model in all the cases is the LSTR model with a single smooth break.

2003 to 2006, is the transition period from the high to the low violence regime that indicates presence of a smooth break. The transition functions for the series depicting the number of casualties have been plotted in Figure 5. The Ljung-box statistics for the residuals of the series show that there is no remaining serial correlation at 1% level of significance and the model specification is appropriate.

Recall that the series for incidents involving explosives is truncated as data is only available from 2001 onwards. Nonetheless, results from the nonlinear model are similar: the high violence regime with incidents involving explosives stretches until mid-2003 and the series transitions up to early 2009 to the low violence state. Detailed graphs for all the five series are given in Figure 6 (Figures 6a to 6e).

Table 5c provides the break dates for the time series for casualties as well as explosion incidents. The dates for all the four series on casualties are clustered in 2005 from January to April, while the break date for the incidents involving explosives is June 2006.²³

Table 5d provides the descriptive statistics of the violence series across the high, transition and low violence regimes. We observe that there has been a secular decline in violence in the state of Jammu and Kashmir over the period of study. This is depicted by the decline in the maximum, minimum and average levels of violence as we transit across the regimes.

4.2 Bai and Perron (BP) Breaks

Enders (2015) suggests that the Bai and Perron (1998, 2003) procedure precludes the possibility of a nonlinear break in the series. We nevertheless use the procedure in the same manner as Enders and Sandler (2005) and Amara (2012) to detect breakpoints as structural breaks in the evolution of a conflict. In this sub-section, we first discuss the exact break dates found by BP procedure and the events that coincide with these breaks. Thereafter, we relate these breaks to the smooth breaks that we detected and discuss the overall evidence.

Table 6 displays results of the BP procedure, i.e. results from estimating equation 6, where the number of breaks (m) is selected by using BIC. Table 6 reports the point estimate

²³ Our results for nonlinear break dates were robust across grid search procedures and different starting values.

for each break date, the 95 percent confidence intervals around the break dates (the columns Lower and Upper), as well as the estimated intercept and trend in the $(m + 1)$ regimes. Hence, on the row for each series, the last two columns report the mean and trend *prior* to the break date for the series.

We find four break dates in total killed and terrorists killed, three each for civilians and incidents involving explosions, two for security force personnel killed, and one for foreign terrorists killed. The series total killed is the sum of terrorists, civilians and security force personnel killed and, thus, reflects the overall situation. As the numbers for the terrorist killed series are by far the highest, the break dates we observe in the series for total killed are similar to those in the terrorist killed series.

The first break-point, identified in all the casualties series, is 2001 which marks the peak in the trend of casualties. The events explaining this break-point are related to, on one hand, diminished support to militant groups from Pakistan in the face of international scrutiny post 9/11 and, on the other hand, Indian security forces regaining control of urban areas lost due to troop movements during the Kargil war. A break-point in 1999 in the civilians killed series is also a possible result of the turbulence during that time period.

In the aggregate casualties series, we find the key break date to be March 2005. This comes shortly after the completion of the fencing of the LoC, indicating that the fence had a significant effect on the level of violence in the state. The break date also happens to be in the same month when the reconstruction plan was implemented although it was announced much earlier. The terrorists killed and foreign terrorists killed also capture the same break-point, as does the incidents involving explosions series. Particularly the drop in the number of foreign terrorists killed close to the completion of the fence highlights the importance of the fence in reducing the influx of terrorists.

Interestingly, we find that the break-points for the civilian and security force personnel casualties series lag behind the break date for the terrorist casualties (2006 November for security force personnel and August 2007 for civilians, respectively). This indicates that the improved security environment provided by the fencing of the LoC, allowed for effective

implementation of the large PMRP and the NREGS, which in turn led to a decline in the number of terrorist killed initially and then the civilians and security forces killed. It subsequently paved way for the success of the economic development programs and, hence, a further decline in violence across the board as suggested by the opportunity cost mechanism.

These break-points are succeeded by break-points in November 2006 for the total and security force personnel casualties series, in March 2007 for the terrorists killed, and in August 2007 for the civilians killed and incidents involving explosions. These break-points are succeeded by a period where violence was lower and declining at a slower pace, especially in the incidents involving explosions series.

Finally, we also observe a break-point in early 2003 for the terrorist and total killed. This break-point is characterized by decrease in violence, and it coincides with the year when India and Pakistan restored their diplomatic ties. This is also the year when the smooth transition of the LSTR model into the low-violence regime starts for all of the series.

Figures 6a to 6f depict the break-points obtained from BP and LSTR models. It is interesting to note that violence begins to decline from 2003 and transits to a low violence regime by the beginning of 2007 in all the cases. This is also the time period where the sharp BP and smooth breaks are clustered. Taken together, the evidence indicates that events taking place between 2003-2006 must have contributed to the reduction in the violence level. This period marks a change with the Indian troops recapturing interior areas post Kargil and the diplomatic dialogue between India and Pakistan resuming in 2003, fencing of the J&K border with Pakistan being completed in 2004, the PMRP coming into force in 2005 and, finally, NREGS being rolled out in 2006. Similarly, the BP and LSTR break dates for the incidents involving explosions are identified between late 2002 and 2008. It seems that the nonlinear methodology captures the transition periods in the series while the BP technique re-emphasizes similar results by highlighting some of the key turning points in the series.

4.3 Robustness checks

We conduct a number of checks to assess the robustness of our findings. First, we check if the results are affected by seasonality. For instance, violence in the state is typically lower in the winter months. Therefore as a robustness check we re-estimate the LSTR models with seasonality adjusted series that obtained using the X-12-ARIMA adjustment. The robustness check of our main specification, the LSTR model, is presented in Table A1 in the online appendix. The results for the deseasonalized data in Table A1 Panel (a) are similar to those given in Table 5b. Compared to Table 5c, the break dates in Table A1b shift ahead by one month for the total, terrorist, and civilians casualties series, while it stays the same for security personnel casualties. However, in the case of the explosions series, the break date shifts back by about three months to March, 2006. Nonetheless, it should be noted that the break dates continue to fall in the same years, that is, 2005-06.²⁴

Second, so far in the analysis we have utilized the violence count data without imposing a Poisson process on the time series models. An alternative approach would be to consider a linear mapping of the count data which maintains the order of the realizations in the original time series sequence. The transformed data constructed in such a way would no longer have counts. As a robustness check, we implement the LSTR model on such transformed data, which we calculate by standardization of the data for violence in the state. The results of the LSTR models for the standardized series are given in Table A2 in the online appendix and indicate that the break dates are the same as those from the LSTR models for the non-standardized count data discussed above.

As a third robustness check we estimate equation 8, which extends equation 7 by adding AR terms as exogenous regressors (Enders and Sandler, 2005). The results are reported in Table A3 of the online appendix and are generally similar to those reported in Table 6. While we find a few more break dates (four for civilians and three for security force personnel), the confidence intervals around the break dates are now narrower, hence, providing firmer evidence in favor of the break dates presented in Table 6. For instance, the confidence

²⁴The results for the Bai & Perron estimation remain similar to the ones presented in the paper and are available from the authors on request.

intervals of the 2005 breakpoints are just four months for the total casualties series, and three for the terrorists series. We also find a break point in April 2005 for the civilian casualties series, confirming our previous findings. Finally, the new break-point for security force personnel killed in 2013 marks the time when the series converges to nearly zero.

Lastly, we also conduct Chow’s (1960) test for the joint significance of the break dates detected using the Bai and Perron procedure and reject the null of no breaks at the 1% significance level in all the series.²⁵

5 Conclusion

Over two and a half decades since the beginning of the conflict, the Indian government continues to search for policies to address the ongoing insurgency in Jammu and Kashmir. In this paper, we use a variety of time series techniques to assess the role played by several military, political and economic measures in reducing conflict in the state. The effect of policy interventions on conflict may gradually manifest over time, making it difficult to ex-ante pinpoint break dates in the time series data on violence. In this study, we go beyond the standard tests for sharp structural breaks used in the literature (Bai and Perron, 1998 and 2003; Chow, 1960; Andrews, 1993) by using endogenous nonlinear smooth break tests based on the LSTR model (Teräsvirta, 1994; 1998; Lin and Teräsvirta, 1994) to examine insurgency in the Indian state of Jammu and Kashmir over the period 1998-2014.

The nonlinear LSTR models indicate transition from a high violence state to a low violence state around 2005 corresponding to the fencing of the border between India and Pakistan. Subsequent large scale employment generation and infrastructure development programs implemented in the improved security environment coincide with a further reduced violence particularly that directed against civilians. The results from the Bai and Perron test procedure further validate this finding, and our results are robust to different model specifications and transformations of the data. This pattern in the timing of breaks

²⁵We test for a break in both in the constant and trend of the series. The results are available from the authors on request.

is indicative of the causal factors that may have been at play during the period of declining violence in the state. In particular, the results provide suggestive evidence on the complementary relationship between security and development programs, which is supported by recent literature.

As further data becomes available, future research could extend the analysis to a more micro level. As discussed earlier, Berman et al. (2011b) find that small-scale projects implemented with local collaboration were successful in reducing violence in Iraq. While most the development projects in India are rather large scale, spatial and temporal variation in their effect on violence and their interaction with security policies could provide a fruitful avenue to gain a deeper understanding of the effect of development programs on conflict.

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Figures and Tables

Figure 1: Map of Kashmir & Jammu



Figure 2: Timeline of Events

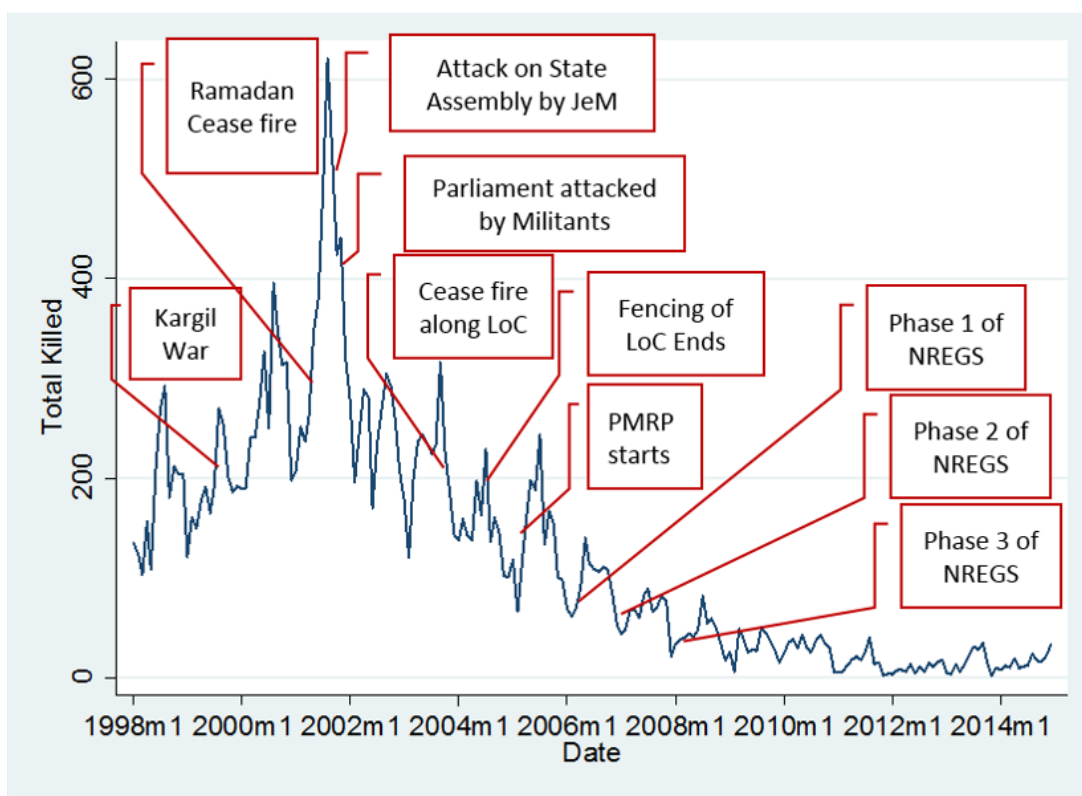
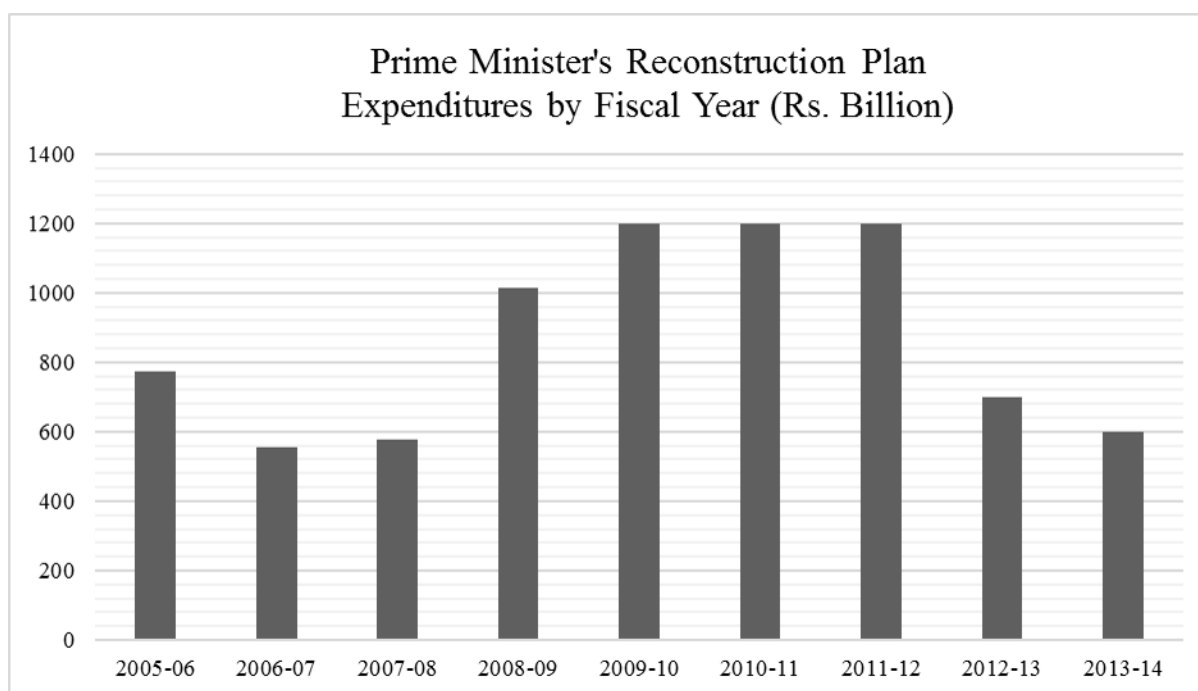
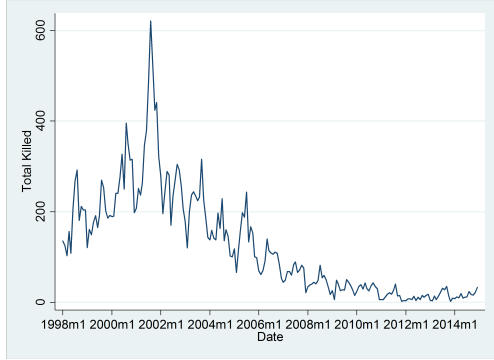


Figure 3: Expenditures of the Prime Minister's Reconstruction Plan

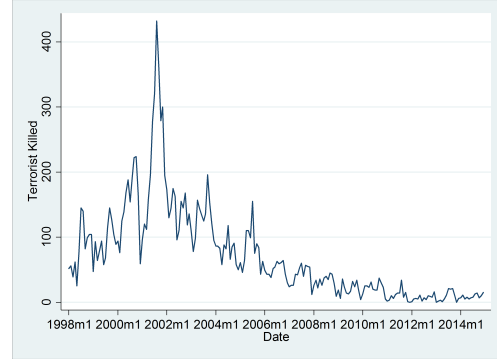


Source: Authors' calculations based on information received from the Ministry of Home Affairs, Government of India. All figures are in nominal terms.

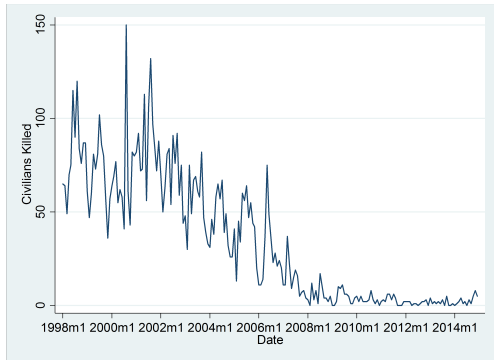
Figure 4: Time Series for casualties in Jammu and Kashmir



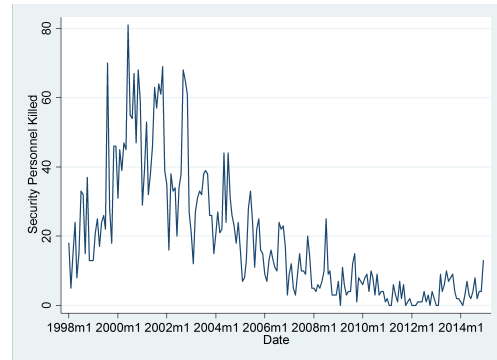
(a) Total Killed



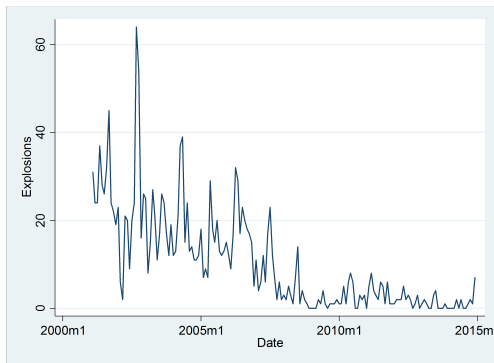
(b) Terrorists Killed



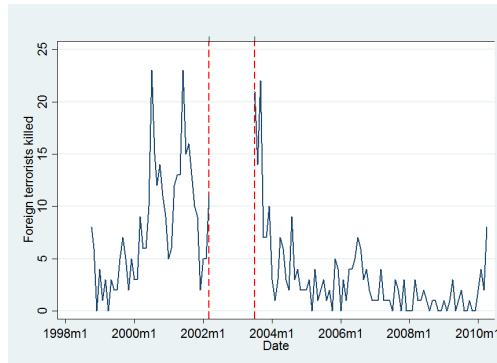
(c) Civilians Killed



(d) Security Personnel Killed



(e) Number of Incidents Involving Explosives



(f) Foreign terrorists killed

Figure 5: Transition Functions Estimated from the LSTR Model

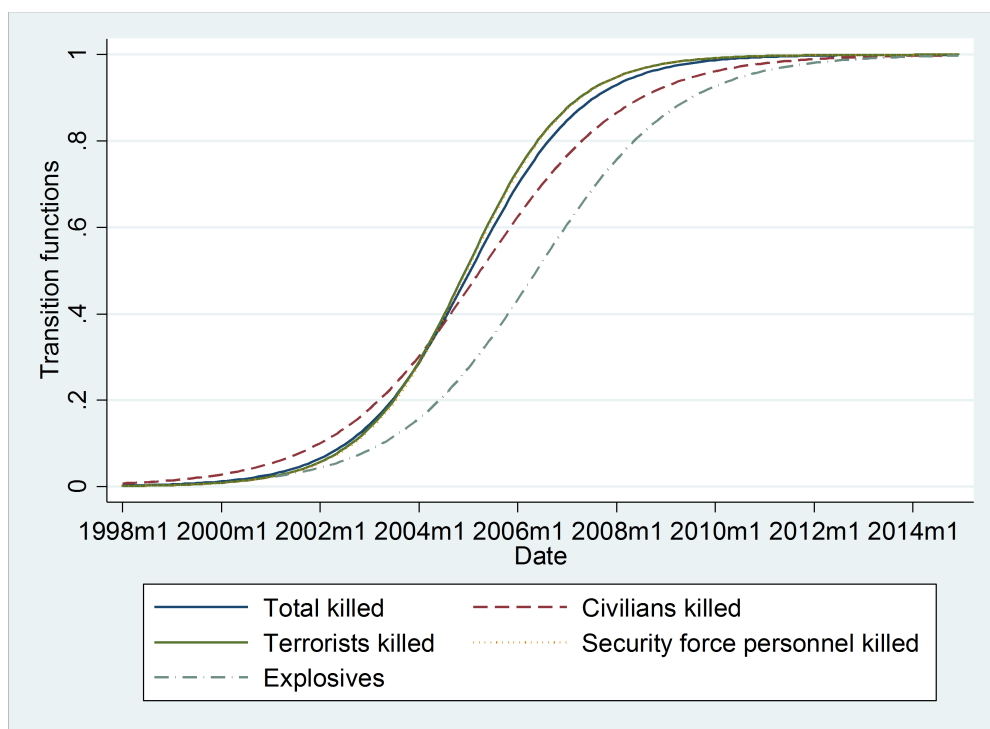
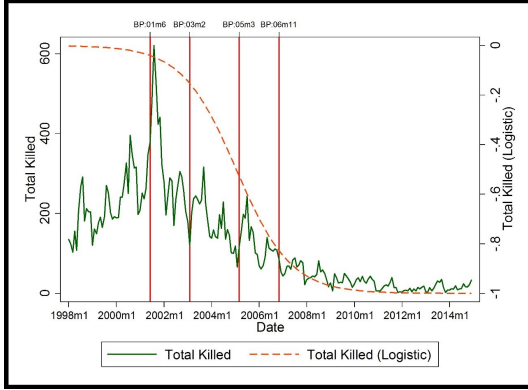
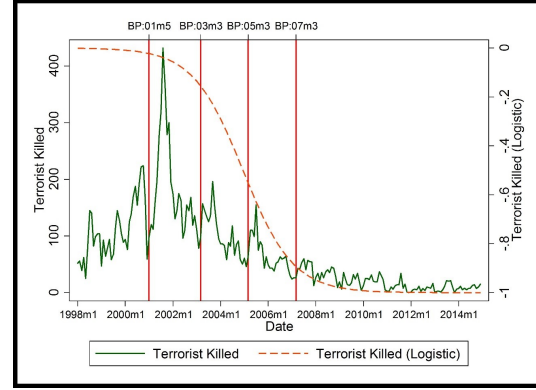


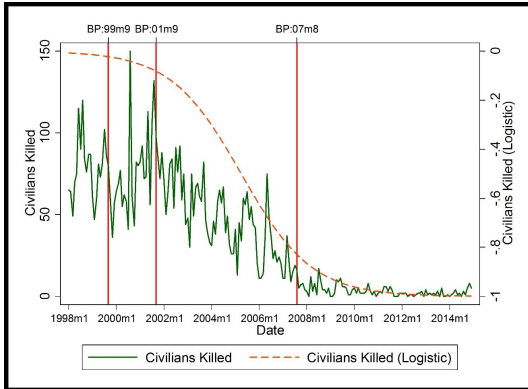
Figure 6: Structural Breaks



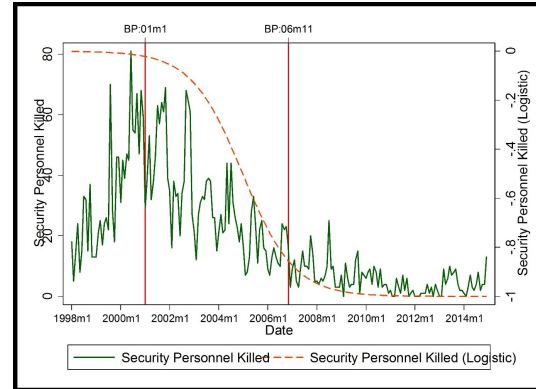
(a) Total Killed



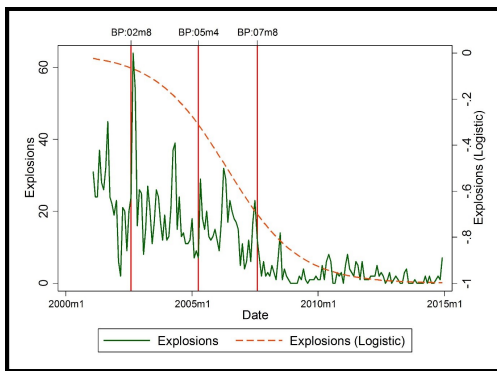
(b) Terrorists Killed



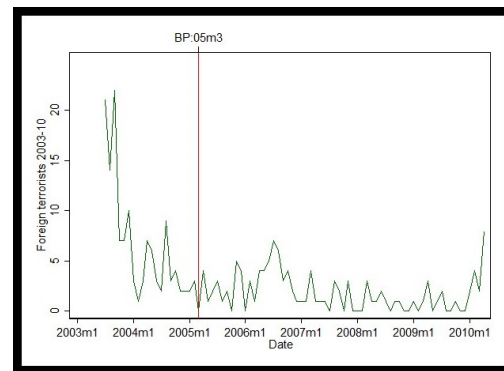
(c) Civilians Killed



(d) Security Personnel Killed



(e) Number of Incidents Involving Explosives



(f) Foreign terrorists killed

Table 1: List of Events

Date	Event
May-July 1999	Kargil War
Dec 2000- May 2001	Ramadan Cease fire
1st Oct 2001	Suicide attack by the JaisheMohammed (JeM) on the State Legislative Assembly complex in Srinagar
13th Dec 2001	Indian National Parliament attacked by Militants
Sept-Oct 2002	State elections, Jammu Kashmir Peoples Democratic Party (PDP) comes to power
25 th Nov 2003	India and Pakistan agree to a cease-fire along the LoC
Sept 2004	LoC fencing completed
Nov 2004	PM's Reconstruction Plan (PMRP) announced.
April 2005	PMRP starts
Feb 2006	Phase 1 of NREGS
April 2007	Phase 2 of NREGS
April 2008	Phase 3 of NREGS

Table 2: Descriptive Statistics

	y_{TOT}	y_{CIV}	y_{TER}	y_{SP}	y_{EXP}	y_{FT}
Mean	123.6225	34.5049	70.2598	18.85784	10.5298	2.99
Std. Dev	116.1502	34.3978	71.1165	18.2096	11.6906	3.94
Skewness	1.1881	0.7268	1.8292	1.2211	1.5694	3.00
Kurtosis	4.5189	2.5820	7.6228	3.8229	5.9879	11.17
Maximum	621	150	432	81	64	22
Minimum	2	0	0	0	0	0
Observations	204	204	204	204	168	82

Note: y_{TOT} denotes the total killed, y_{CIV} denotes the civilians killed, y_{TER} denotes the terrorist killed, y_{SP} denotes the security personnel killed, y_{EXP} denotes the incidents involving explosives and y_{FT} foreign terrorists killed.

Table 3: Lee-Strazicich Test for a Unit Root with Structural Breaks

Variable	Trend Break Model	Inference
y_{TOT}	-59.355	I(0)
y_{CIV}	-64.992	I(0)
y_{TER}	-75.268	I(0)
y_{SP}	-57.886	I(0)
y_{EXP}	-6.069	I(0)
y_{FT}	-4.588	I(0)

Critical Values				
Trend Break Model		λ_2		
λ_1	0.4	0.6	0.8	
0.2	-6.16, -5.59, -5.27	-6.41, -5.74, -5.32	-6.33, -5.71, -5.33	
0.4	-	-6.45, -5.67, -5.31	-6.42, -5.65, -5.32	
0.6	-	-	-6.32, -5.73, -5.32	

Note: y_{TOT} denotes the total killed, y_{CIV} denotes the civilians killed, y_{TER} denotes the terrorist killed, y_{SP} denotes the security personnel killed and y_{EXP} denotes the incidents involving explosives and y_{FT} denotes foreign terrorists killed. Critical values are at the 1%, 5% and 10% levels, respectively. λ_j denotes the location of breaks.

Table 4: Tests for Nonlinearity

Variable	Tsay (1986) Test Statistic	Luukkonen et al. (1988) Test Statistic
y_{TOT}	5.59***	3.42***
y_{TER}	4.55***	4.26***
y_{CIV}	8.19***	5.43***
y_{SP}	6.10***	3.95***
y_{EXP}	4.33***	2.96***

Note: y_{TOT} denotes the total killed, y_{CIV} denotes the civilians killed, y_{TER} denotes the terrorist killed, y_{SP} denotes the security personnel killed and y_{EXP} denotes the incidents involving explosives. *, **, *** denote significance at 10%, 5% and 1% respectively.

Table 5: Estimates of Smooth Transition Regressions

(a) Auxiliary Regressions to Test for Time as the Threshold Variable

Variable	F-statistic	p-value
y_{TOT}	6.92	0
y_{TER}	5.08	0.002
y_{CIV}	29.32	0
y_{SP}	10.05	0
y_{EXP}	7.95	0

(b) Estimates of LSTR model with time as the threshold variable

Variable	α_0	α_1	α_2	λ_0	c	γ	LB stat
y_{TOT}	58.044***	0.780***	-	-53.366***	85.477***	0.073*	4.212 (0.5193)
y_{TER}	28.127***	0.810***	-	-25.279***	84.320***	0.079	2.029 (0.8451)
y_{CIV}	58.220***	0.263***	-	-57.500***	87.940***	0.056***	2.236 (0.8156)
y_{SP}	15.921***	0.456***	0.141**	-13.972***	84.71***	0.079*	6.189 (0.2882)
y_{EXP}	14.310***	0.415***	-	-13.632***	65.609***	0.058**	10.368 (0.0655)

Table 5: Estimates of Smooth Transition Regressions (cont.)

(c) Estimated Break Dates

Variable	c	Break Date
y_{TOT}	85.48***	February, 2005
y_{TER}	84.32***	January, 2005
y_{CIV}	87.94***	April, 2005
y_{SP}	84.71***	January, 2005
y_{EXP}	65.61***	June, 2006

(d) Descriptive Statistics across Regimes

Variables	High-Killing Regime			Transition Phase			Low-Killing Regime		
	Mean	Max	Min	Mean	Max	Min	Mean	Max	Min
y_{TOT}	253.5	621	103	148.6	316	44	25.2	82	2
y_{TER}	77.3	150	36	39.4	92	0	2.8	11	0
y_{CIV}	140.5	432	25	84.6	196	24	18.0	60	0
y_{SP}	37.3	81	5	22.1	65	3	5.3	25	0
y_{EXP}	25.5	64	2	12.0	39	0	2.1	8	0

Note: y_{TOT} denotes the total killed, y_{CIV} denotes the civilians killed, y_{TER} denotes the terrorist killed, y_{SP} denotes the security personnel killed and y_{EXP} denotes the incidents involving explosives. The optimal number of lags included in the model are one for civilians killed, terrorist killed and total killed, two for security personnel killed and one for incidents involving explosives. *, **, *** denote significance at 10%, 5% and 1% respectively. Column LB stat in (b) denotes the Ljung-Box test statistic for autocorrelation in the residuals.

Table 6: Estimates of Bai and Perron Multiple Structural Breaks

Series	Breakdate	Lower	Upper	Intercept	Trend
y_{TOT}	2001:06	2000:12	2001:07	143	3.99
	2003:02	2003:01	2005:06	1213	-17.6
	2005:03	2004:11	2005:08	629	-6.15
	2006:11	2006:10	2007:07	620	-5.13
y_{TER}				114	-0.55
	2001:05	2000:10	2001:06	58	2.65
	2003:03	2003:02	2004:08	819	-12.22
	2005:03	2005:01	2005:07	440	-4.57
y_{CIV}	2007:03	2007:02	2008:03	411	-3.53
				81	-0.4
	1999:09	1999:07	2000:01	75	0.38
	2001:09	2001:08	2001:11	6	2.15
y_{SP}	2007:08	2007:07	2007:11	113	-0.83
				11	-0.05
	2001:11	2001:10	2002:01	13	1.08
	2006:11	2006:10	2008:01	59	-0.44
y_{EXP}				15	-0.06
	2002:08	2002:05	2003:01	35	-1.17
	2005:04	2004:12	2008:01	45	-0.70
	2007:08	2007:07	2008:04	34	-0.27
y_{FT}				6	-0.03
	2005:03	2005:02	2007:01	60	-0.71
				5	-0.03

Note: y_{TOT} denotes the total killed, y_{CIV} denotes the civilians killed, y_{TER} denotes the terrorist killed, y_{SP} denotes the security personnel killed, y_{EXP} denotes incidents involving explosives and y_{FT} denotes foreign terrorists killed. Shifting regressors are constant and a trend. The lower and upper columns denote 95% confidence intervals. The intercept and trend are the coefficient estimated for the regime prior that break. Number of breaks is determined by BIC with the maximum number of breaks set to 5 and the minimum length of the regime set to 20, except in the case of y_{FT} where the maximum number of breaks set to 3 and the minimum length of the regime set to 15 due the shorter time series.

Online Appendix for:
Do Fences Make Good Neighbors? Evidence from an
Insurgency in India

Kaila, Singhal, Tuteja

A1 LSTR Results with Deseasonalized Data

Table A1: LSTR Model for the Transformed Data

(a) Estimates of LSTR model with time as the threshold variable

Variable	α_0	α_1	α_2	λ_0	c	γ
y_{TOTD}	58.741***	0.789***	-	-50.253***	85.882***	0.074**
y_{TERD}	27.549***	0.810***	-	-24.675***	85.236***	0,081
y_{CIVD}	70.806***	0,103	-	-69.929***	88.205***	0.056***
y_{SPD}	15.819***	0.384***	0.213***	-13.869***	84.824***	0.080**
y_{EXPD}	14.211***	0.444***	-	-13.619***	63.242***	0.053**

(b) Estimated Break Dates

Variable	c	Break Date
y_{TOTD}	85.88***	March, 2005
y_{TERD}	85.23***	February, 2005
y_{CIVD}	88.21***	May, 2005
y_{SPD}	84.82***	January, 2005
y_{EXPD}	65.61***	March, 2006

Note: y_{TOTD} denotes the total killed, y_{CIVD} denotes the civilians killed, y_{TERD} denotes the terrorist killed, y_{SPD} denotes the security personnel killed and y_{EXPD} denotes the incidents involving explosives using the deseasonalized data. The optimal number of lags included in the model are one for civilians killed, terrorist killed and total killed, two for security personnel killed and one for incidents involving explosives. *, **, *** denote significance at 10%, 5% and 1% respectively. Column LB stat in (b) denotes the Ljung-Box test statistic for autocorrelation in the residuals.

A2 Nonlinear Breaks in Transformed Data

As a robustness check, we transform our data using a linear mapping which preserves the order of the values. This is achieved by standardization of the time series by deducting the mean and dividing by the sample standard deviation. For a series x_t , then, the transformed series will be given by $x_{tT} = \frac{x_t - \text{mean}}{\text{standard deviation}}$. In order to confirm the break dates, we redo the nonlinear analysis on the transformed data and estimate the LSTR model to detect the smooth breaks. The results are given in Table A2 below.

From Table A2 we find that the results for the transformed data are identical to those stated before. The break dates are correspondingly the same as before. This suggests that the break dates are robust.

Table A2: LSTR Model for the Transformed Data

(a) Estimates of LSTR model with time as the threshold variable

Variable	α_0	α_1	α_2	λ_0	c	γ
y_{TOTT}	0.266***	0.780***	-	-0.459***	85.478***	0.073*
y_{TERT}	0.209***	0.811***	-	-0.356***	84.320***	0.079
y_{CIVT}	0.953***	0.263***	-	-1.672***	87.940***	0.056***
y_{SPT}	0.457***	0.456***	0.141**	-0.768***	84.717***	0.079*
y_{EXPT}	0.697***	0.415***	-	-1.166***	65.609***	0.058**

(b) Estimated Break Dates

Variable	c	Break Date
y_{TOTT}	85.48***	February, 2005
y_{TERT}	84.32***	January, 2005
y_{CIVT}	87.94***	April, 2005
y_{SPT}	84.72***	January, 2005
y_{EXPT}	65.61***	June, 2006

Note: y_{TOTT} denotes the total killed, y_{CIVT} denotes the civilians killed, y_{TERT} denotes the terrorist killed, y_{SPT} denotes the security personnel killed and y_{EXPT} denotes the incidents involving explosives using the transformed data. The optimal number of lags included in the model are one for civilians killed, terrorist killed and total killed, two for security personnel killed and one for incidents involving explosives. *, **, *** denote significance at 10%, 5% and 1% respectively. Column LB stat in (b) denotes the Ljung-Box test statistic for autocorrelation in the residuals.

A3 Bai & Perron Results with Exogenous AR-terms

Table A3: Estimates of Bai and Perron Multiple Structural Breaks (with exogenous AR terms)

Series	Breakdate	Lower	Upper	Intercept	Trend
y_{TOT}	2001:06	2001:05	2001:07	150	5.89
	2003:02	2003:01	2003:05	1227	-16.12
	2005:03	2005:02	2005:05	751	-6.96
	2007:11	2007:10	2007:12	563	-4.18
				128	-0.6
y_{TER}	2001:05	2001:04	2001:06	46	3.38
	2003:03	2003:02	2003:06	680	-9.2
	2005:02	2005:01	2005:03	446	-4.47
	2006:10	2006:09	2006:11	332	-2.66
				90	-0.44
y_{CIV}	2001:04	2001:02	2001:05	98	0.14
	2003:02	2003:01	2003:12	265	-3.05
	2005:04	2004:11	2005:06	183	-1.54
	2007:04	2007:03	2007:06	213	-1.65
				20	-0.1
y_{SP}	2001:11	2001:10	2000:12	-14	3.6
	2006:07	2006:06	2006:09	137	-1.12
	2013:02	2013:01	2014:02	57	-0.31
				23	-0.07
y_{EXP}	2003:08	2003:07	2003:10	37	0.52
	2006:02	2006:01	2006:04	55	-0.45
	2008:07	2008:06	2008:08	116	-1.18
				10	-0.04

Note: y_{TOT} denotes the total killed, y_{CIV} denotes the civilians killed, y_{TER} denotes the terrorist killed, y_{SP} denotes the security personnel killed and y_{EXP} denotes incidents involving explosives. Shifting regressors are constant and a trend. The lower and upper columns denote 95% confidence intervals. The intercept and trend are the coefficient estimated for the regime prior that break. Number of breaks is determined by BIC with the maximum number of breaks set to 5 and the minimum length of the regime set to 20. The AR structure is chosen by using AIC, BIC and HQ criteria. The number of exogenous AR-terms chosen are 8, 9, 12, 12 and 12 for the series Total killed, Terrorists, Civilians, Security Force Personnel, and Explosions, respectively.

A4 The Fence

Figure A1: The fence at the Line of Control (source: Rediff news)

