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Employment in Spanish regions: Cost control or growth-enhancing policies?*

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Abstract

Spain provides an extreme case of unemployment rate oscillations (8.3% in 2007, 26.1% in 2013, 19.6% in 2016) in parallel with acute regional persistence in labour market outcomes –the sets of relatively high and low unemployment regions have not changed in decades. Since generic labour market reforms have been fruitless on this respect, we explore whether such groups of regions react differently to key drivers of employment and wage setting. We find that the low income (high unemployment) regions are more reactive to capital accumulation, and thus to a growth strategy based on stimulating investment. In turn, the high income (low unemployment) ones are more sensitive to the wage-productivity gap, and thus to a strategy that keeps unit labour costs (ULC) low. Such patterns call for more region-specific policies and discard standard labour market reforms as a unique tool to manage the unemployment rate problem. Further, to the extent that investment serves both at fostering capital accumulation and labour productivity (which, in turn, reduces the ULC), regionally-targeted soft credit lines and capital taxes could be helpful in breaking regional sluggishness.

JEL Classifications: R11, E24, E22, J23, J31.

Keywords: Employment, Wage setting, Labour income share, Capital stock.

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1 Introduction

One of the main expressions of the Great Financial Crisis (GFC) was unemployment, with Spain reaching a 26.1% in 2013, the second highest rate, after Greece, among the OECD countries. In turn, the recovery has been characterised by the low quality of new jobs with massive use of non-permanent contracts and lower-than-before starting wages. This is largely the consequence of the austerity and internal devaluation policies required to achieve the external adjustment.

These policies, in particular the labour market reforms, are common across regions (and to some extent even across countries) notwithstanding the fact that the labour market situation of such regions may be structurally different. For example, in 2013, Andalusia was the Spanish region with the highest unemployment rate (36.2%), while the Basque Country recorded the minimum one (16.6%). This huge contrast is not the outcome of the crisis, but a structural situation since the 1980s, when Spanish unemployment started hovering around 20%.

In view of this persistent asymmetry, could one envisage more targeted policies than generic labour market reforms? To respond to this question, we seek to explore to what extent employment reacts differently, across regions, to key labour demand determinants. Discrepant reactions would call for region-specific policies that should be regarded as complementary to the standard generic labour market reforms.

Of course, generic labour market policies have different effects across regions on account of the diverse composition of employment in many dimensions (age, gender, education, industry composition, or the duration of unemployment). However, the Spanish experience shows an extreme regional unemployment persistence with equal sets of relatively high and low unemployment regions today than in the 1980s. It thus seems that the time has come for more region-specific policies.

From a theoretical perspective, the labour demand is usually derived from a production function featuring labour, capital, and technology as production factors. Hence, in any standard setting, wages (as the price of labour), capital stock, and some measure of efficiency appear as key determinants of labour demand decisions. Here we consider labour productivity as the measure of efficiency, and focus on the role exerted by the labour share as determinant of employment.¹ We regard this as an important contribution for a twofold reason. First, there is growing literature on the causes behind the fall in the labour share since the 1980s (see, for example, Karabarbounis and Neiman, 2014), but not much on its consequences. Thus, we provide new evidence on its employment

¹The labour share is the total compensation received by the labour factor over output (Y). Given that total compensation can also be seen as the average compensation per employee (W) times the number of employees (N), we have: $\frac{WN}{Y} = \frac{W}{Y/N}$. Empirically, we will use logs and deal with $w_t - pr_t$, where $pr_t = y_t - n_t$ (lowercase letters denote natural logs of the corresponding capital letters).

effects. Second, in focusing on the labour share, we are actually looking at the effects of the wage-productivity gap (since a fall in the labour share is equivalent to a situation in which wages grow by less than labour productivity). This implies that we can check the relative incidence on employment of (i) decreases in the labour share (and thus increases in the capital share) –which can be associated to the internal devaluation strategy followed to rebalance the economies after the GFC–, versus (ii) the enlargement of capital stock –which can be associated to the alternative focus on growth policies, which have also been defended as a possible way out strategy. Of course, we do not claim that we capture the effects of imaginary policies. We just argue that our analysis can yield insights into the effects of some regional specific action connected to either of these two policy strategies.

This is relevant because the widening of the wage-productivity gap has been one of the pillars behind the adjustment in the aftermath of the GFC (it was ensured through labour market reforms to tie up wages, at the same time that accelerated job destruction relative to output fall enhanced labour productivity). The higher the wage-productivity gap, the lower the unit labour cost, and the greater the expected price-competitiveness gains that can be achieved to restore external imbalances. This internal devaluation policy, however, has been criticized (for example, by Paul Krugman or Joseph Stiglitz) and sometimes judged as self-defeating when implemented together with austerity. Against this view, the proposed alternative has been to focus on growth-enhancing policies aiming, for example, to reboot investment. Along this line, our analysis provides new evidence on the impact that the intensive growth in capital accumulation had on job creation during the expansion and, conversely, on job destruction when investment rates faded along the GFC. This is done at the regional level in search of specificities that could bring new advice in the design of the policy strategy.

We conduct this analysis for the Spanish economy because it is probably the most extreme case of unemployment volatility displaying, simultaneously, a striking and persistent contrast among regions regarding their labour market performance. Volatility in the sense that it has experienced acute convergence and divergence processes with respect to the European average across business cycle phases. Indeed, the unemployment rate quickly converged to the European average during the ‘wild ride’, evolving from 24.2% in 1994 to 8.2% in 2007, and then diverged much intensively to reach 26.1% in 2013. From a regional perspective, this was expressed in the form of major swings in regional unemployment rates, but still high persistence in the relative ranking of regions.

We exploit the contrast in regional economic performance to cluster the 17 Spanish regions in two well-defined groups. To avoid endogeneity, instead of splitting these regions according to their employment dynamics, we choose relative income as an alternative variable.² As Table 1 shows, the first resulting group includes regions with relatively

²Relative regional income is defined as income in region i divided by national income. Thus, regions

low income, while the second one is made of the high income regions. This classification fully coincides with the one that would arise from grouping them into high and low unemployment rate areas.³

Table 1. Classification of Spanish regions.

HIGH INCOME REGIONS	LOW INCOME REGIONS
Aragon	Andalusia
Balearic Islands	Asturias
Catalonia	Canary Islands
Madrid	Cantabria
Navarre	Castilla-La Mancha
Basque Country	Castilla and Leon
La Rioja	Extremadura
	Galicia
	Murcia
	Valencia

Note: Regions clustered depending on their relative income.

Aside from studying the determinants of employment per group of regions, we search for wage setting regional specificities. In a standard framework, a one-to-one relationship between wages and labour productivity is to be expected (see, for example, Hatton, 2007). If such relationship holds, the main implication is the absence of a changing wage-productivity gap or, in other words, a constant labour share. If that was the case, our analysis in terms of employment would be largely irrelevant because no employment variation could be attributed to such a constant. If, however, this one-to-one relationship does not hold, there is room to examine the extent to which the wage-productivity gap affects employment. Hence, it is important to assess the long-run elasticity of wages with respect to productivity (of course in the presence of the appropriate controls) and check whether significant differences exist across groups of regions.

The study of the employment and wage setting determinants is conducted through the estimation of labour demand and wage setting equations for both sets of regions in years 1980-2011. As explained in Section 3, for each set of regions and type of equation, dynamic panel data estimations by Ordinary Least Squares (OLS) and the Generalised Method of Moments (GMM) are conducted following the Autoregressive Distributed Lagged (ARDL) approach to cointegration analysis.

where this indicator is, on average, above unity are clustered in a high income group, while the remainder are grouped in the low income group of regions.

³Bande and Karanassou (2009, 2014) use the relative unemployment rate as the classification criteria for the Spanish regions, and find a similar distribution. This justifies the association between high (low) income and low (high) unemployment regions.

Our empirical models show substantial differences in the determinants of employment in the High and Low income groups of regions. In addition, the estimates of wage setting equations reveal that the one-to-one long run correspondence between productivity and wages is not supported by the data, paving the way for a significant impact of the labour share on regional employment. We verify that this impact is different across regions and use our estimates to perform a dynamic accounting exercise in which we explore how much of the variation in employment in each group of regions can be attributed to changes in the labour share, and how much to changes in capital stock, which is the other crucial determinant.

Our analysis shows that the labour share and capital accumulation have jointly played a major role in explaining employment swings in the more developed regions. In contrast, we find that low income regions do not respond to the labour share, being investment the main source of employment variation. More precisely, the labour share explains 18.4% of employment variations in 1993-1999, and 14.2% in 1999-2007 in the high income regions, while these values in the low income regions are only 3.0% and 1.7% respectively. Investment, on the contrary, explains 36.0% and 23.9% of the employment variations in the low income regions during the 1993-1999 and 1999-2007 periods, respectively, while these contributions are of 15.5% and -3.7% in the high income regions. These results support the view that Spanish regions have experienced a dual employment growth pattern.⁴

We also explore the role of productivity on wage setting, and how their dynamic interaction shapes the evolution of the labour share and employment in both groups. In a scenario with no productivity growth, the wage gap would have been reduced in the low income regions, but it would have risen in the high income ones. As a consequence, employment would have evolved below its actual path in the former, and remained essentially unchanged in the latter.

We also find a different employment impact in a scenario in which productivity gains are fully translated into wage growth. As in the previous scenario, the labour share would have increased and caused a fall in employment in the low income group. In the high income group, on the contrary, the labour share would have experienced a smaller fall, but employment would have quite followed its actual path.

The results on this counterfactual analysis can be used to provide some economic policy advice.⁵ If high income regions are more sensitive to changes in unit labour costs,

⁴This result was already hinted by Bande *et al.* (2008), who argued that wage imitation effects in a context of a semi-decentralised wage bargaining system could result in regional differences in unit labour costs, which were consistent with the evolution of regional unemployment disparities. Note that the concept of labour unit cost is intimately related to that of the wage gap.

⁵We are fully aware that the counterfactual analysis cannot be interpreted as evidence on how would the Spanish regional labour markets *actually* have evolved under the assumptions made (i.e., the chosen counterfactual). However, such analysis is enlightening when trying to account for (and learn on) the major drivers of employment in past significant periods.

standard policies aiming at enhancing wage flexibility will have a larger incidence in those regions than in the low income regions. Hence, a side effect in periods such as the current one, in which these type of policies dominate in a context of low investment rates, will be a contribution to unemployment persistence. Given that low income regions benefit to a lesser extent from these generic policies, complementary measures aiming at fostering investment should also be implemented so that these regions are given the chance to catch up in terms of labour market performance. Therefore, regionally-targeted soft credit lines and region-specific capital taxes could be helpful in breaking regional sluggishness. It is not through cost control measures that regions seem to have more chances to catch up in (un)employment terms.

Our simulations show, in addition, that policies focused on the promotion of low productivity activities are deemed to fail since their effects through wages would be employment-neutral in the more developed regions, and would compromise job creation in the low income group. In spite of this, such activities –construction, hostelry– have led economic growth in Spain in the last two decades. Alternative policies aimed at increasing the link between wage and productivity growth would neither help in alleviating the unemployment problem if productivity growth remains stagnant. It is rather a growth-enhancing strategy with a clear region-specific orientation what should guide the design of policy.

The structure of this paper is the following. In Section 2, we revise the theoretical background endorsing our empirical modelling. The data, empirical methodology, and estimated models are presented in Section 3. These results are used in Section 4 to conduct dynamic simulations on counterfactual scenarios and learn on the recent experience of the Spanish regions. Section 5 concludes.

2 Theoretical background and estimated models

2.1 Labour demand

The labour demand function can be modelled departing from cost minimization or profit maximization setups; it can be unconstrained, in which case its determinants are the factor prices (normally wages and the user cost of capital, although the price of intermediate goods may also be considered); it can also be constrained, in which case it may be constrained by holding capital constant (and thus the stock of capital appears as explanatory variable), or by holding output constant (so that it is output, instead, what acts as explanatory variable); factor prices may be considered in nominal terms or in real terms; and the product market from which the labour demand is derived may be modelled as in

perfect or imperfect competition.⁶

Within this wide range of possibilities, authors have followed different routes depending on the type of analysis being conducted and existing data limitations. For example, Slaughter (2001) relies on unconstrained and constant-output-elasticity labour demands, while Hijzen and Swaim (2010) choose to work with a capital-constrained labour demand model in view of the measurement problems they would face in case of having to use the user cost of capital. In contrast, Adam and Moutos (2014) estimate a constrained labour demand holding output constant, not capital. Irrespective of the approach used, the crucial issue underlying the specific modelling of the labour demand is to provide a correct interpretation of the estimated coefficients.⁷

In our case, we work with a capital-constrained labour demand function such that:

$$N_t = f (W_t , R_t , K_t , PR_t), \quad (1)$$

where N denotes employment, W gross average compensation, R real interest rates, K capital stock, and PR labour productivity. Other controls, conditional on data availability, are also included. These relate basically to demand-side and foreign determinants. Demand-side variables such as private consumption (over GDP) or factors affecting private consumption (such as direct and indirect taxes) affect labour market outcomes through a diversity of channels; for example, via the price elasticity of product demand (Lindbeck and Snower, 1994). In turn, foreign determinants may also play a role, as explained by Rodrik (1997) and explored by Krishna *et al.* (2001) and Hijzen and Swaim (2010), among others. For example, Adam and Moutos (2014) use the real effective exchange rate to capture changes in world demand for the industry's product whose labour demand they estimate. In our case, we do not have region-specific external determinants, and we just control for aggregate (at the Spanish level) foreign effects through the degree of trade openness.

It is standard in this literature (see, for example, Antràs, 2004) to assume that technological progress grows at a constant rate. In that case, the proxy for technology (usually denoted as A) is a linear trend which, in a Cobb-Douglas framework, displays exactly the same coefficient than wages with the opposite sign. Our empirical analysis substitutes the standard linear trend by labour productivity. This is important for a twofold reason.

⁶Standard assumptions include that firms' costs are linearly homogeneous in nominal inputs prices; the production function is linearly homogeneous and features constant returns to scale; markets clear; labour and capital services are supplied elastically to the firm; there is perfect competition in the product market (and therefore prices equal their marginal cost), and factor prices are rewarded their marginal productivities.

⁷For example, in a capital-constrained labour demand equation, the estimated wage coefficient is a proxy of the total effect, as defined by Hamermesh (1993), while in a labour demand holding out constant it is just the substitution effect what is captured (that is, the elasticity of substitution between capital and labour).

First, in the context of the Spanish regions, to have a linear trend is highly restrictive since it would imply a complete absence of technological (internal) catching-up, in spite of the policies implemented towards this objective (not only at the national level, but also at the European level through the large inflow of cohesion and structural funds). Second, to have labour productivity not only allows variation in economic efficiency across regions, but also across time.⁸ Finally, if we can verify, empirically, that wages and labour productivity display the same coefficient with opposite sign, we will be able to implement a restriction so that the labour income share (defined as the log of wages minus the log of labour productivity) appears as determinant of employment. This is an interesting variable to examine given its structural tendency to fall since the 1980s worldwide, which is not fully reflected across regions in Spain. It is interesting, in addition, because much of the labour market adjustment to the globalization process, and the external adjustment in the aftermath of the GFC, has been achieved by trying to push down wages and push up productivity. It is therefore crucial to learn how this pressure over the labour share of income is being successful in terms of job creation.

2.2 Wage setting

The wage setting curve is a curve representing a positive relationship between wages and employment (or negative with respect to unemployment). It can be obtained from nominal wage setting and price setting rules (Nickell, 1998; Nunziata, 2005), but it can also be conceived in terms of real wages (Hatton, 2007).

Two major strands of literature justify the positive relationship between wages and employment. The first one is related to efficiency wages, the second one to the insider-outsider models. The existence of imperfect information in labour markets gives rise to moral hazard and adverse selection problems. In order to minimise the relevance of such problems, firms have incentives to set wages above market clearing levels because such higher wages allow a better control of the workers and a more efficient selection process. The reason is that unemployment resulting from higher than full employment wages: (i) acts as a worker self-discipline device, and thus reduces the moral hazard problem (Shapiro, and Stiglitz, 1984); and (ii) allows for a tighter filtering of candidates thereby curtailing the adverse selection problem.

In the case of Insider-Outsider models (Lindbeck and Snower, 2001) the workers take advantage of existing labour market regulations to push for higher wages. Firms have to face labour turnover costs (LTC) on account of these regulations, and are willing to accept higher than equilibrium wages to avoid incurring in such costs. For a given level

⁸Adam and Moutos (2014) have also used labour productivity in their estimated labour demand models. However, their aim is “to control for shifts in labour demand arising from changes in firms’ required labour to produce a given level of output, as well as in firms’ ability to pay for labour services”.

of LTC, the higher the employment levels, the more strong is the workers' position and the more they press to secure a better compensation.

The idea of effort or enhanced productivity surrounding the efficiency wage models, in their moral hazard or adverse selection versions, connects wages with labour productivity in the associated microfounded models. In turn, the idea of labour turnover costs and, more generally, the incidence of labour market regulations and institutions on labour costs and the wage bargaining position of workers, is what brings into the scene the so-called wage pressure factors.

Blanchard and Katz (1999) reconciled these two strands of literature, which are to be seen as complementary, by postulating a benchmark wage setting equation such that:

$$W_t = f (PR_t , u_t , wage - push\ elements) , \quad (2)$$

where PR denotes labour productivity, u is the unemployment rate, and the wage-push factors are representative of four wide areas of regulation: (i) Union power; (ii) Employment protection legislation; (iii) Unemployment protection legislation; and (iv) Fiscal wedges.

There is abundant literature on the role of such wage-pressure factors on unemployment, which is of course conditioned on data availability for the economy under scrutiny. In our case, we consider two relevant elements. The real minimum wage, which is a regulation protecting employees and has been operational during all our sample period; and social security benefits (as a percent of GDP), which is representative of the overall welfare state, and we regard as a summary variable of the regulations affecting workers (note that such benefits are the counterpart of payroll taxes and, simultaneously, reflect all protecting legislation channelled via the Social Security). This is in addition to oil prices, considered on account of the pressure they are likely to exert on prices, which could have a detrimental impact on real compensation (hence the expected negative sign).

A key elasticity amid the wage setting curve is the one measuring how sensitive wages are with respect to labour productivity. Blanchard and Katz (1999) state that real wages and productivity should be homogeneous of degree 1 in the long run because this is consistent with the fact that technological change does not lead to a persistent trend in unemployment. However, the fact that growth drivers in general, and technological change in particular, have no persistent effects on unemployment has been contested in the literature (see, for example, Karanassou *et al.*, 2010, and references therein).

In our study we leave this issue as an empirical matter. We do not a-priori constrain the long run elasticity of wages with respect to labour productivity because this could conceal relevant contrasts in regional behaviour. Moreover, in case this one-to-one relationship does not hold, it would imply the existence of a persistent wage-productivity gap (since wages do not catch up with productivity even in the long run). We explore

the consequences of such situation on employment because most of the labour market policies implemented in Spain have tended to focus on controlling labour costs, mainly (in successive waves of labour market reforms) through a universal use of temporary jobs, a generous system of payroll tax allowances on new permanent contracts, and cuts in the actual cost of severance payments. In contrast, there has been little emphasis in policies aiming to foster labour productivity as indicated (in labour market terms) by the low development of active labour market policies, and the low incidence of on-the-job training, which has not been targeted by public administrators. As a result, labour productivity in Spain behaves in a complete counter-cyclical manner since the rise and fall in employment, in good and bad times respectively, always exceeds the rise and fall in output.

How does the wage-productivity gap affect employment across types of regions? How has the emphasis of recent implemented policies on internal devaluation affected job creation across regions? Would fostering growth and gross capital formation be an alternative? Where and to what extent? These are questions we seek to answer with the analysis that follows.

2.3 Estimated models

Following the previous discussion, we estimate the following dynamic models.

The labour demand model is specified as:

$$n_{i,t} = v_i^N + \mu_t^N + \sum_{j=1}^J \alpha_j^N n_{i,t-j} + \sum_{j=0}^J \beta_j^N ls_{i,t-j} + \sum_{j=0}^J \gamma_j^N k_{i,t-j} + \sum_{j=0}^J \delta_{s,i,j}^N \mathbf{X}_{s,i,t-j}^N + \varepsilon_{i,t}^N, \quad (3)$$

where n is employment, ls is the labour income share, k is the stock of capital, and \mathbf{X} is a vector of s additional controls; t denotes time, i denotes region, and j denotes lags; v_i represent regional dummies, μ_t represent time dummies; the α 's, the β 's, the γ 's and the δ 's are parameters to be estimated; and ε is the error term.

In turn, the wage setting model is represented by:

$$w_{i,t} = v_i^W + \mu_t^W + \sum_{j=1}^J \alpha_j^W w_{i,t-j} + \sum_{j=0}^J \beta_j^W pr_{i,t-j} + \sum_{j=0}^J \delta_{s,i,j}^W \mathbf{X}_{s,i,t-j}^W + \varepsilon_{i,t}^W, \quad (4)$$

where w is the gross real wage, and pr is labour productivity.

3 Results

3.1 Data

Our data sources are diverse and include the OECD's *Economic Outlook*, the BD-REMS dataset (from the Spanish Ministry of Finance), and Datastream. All regional data are taken from the BD-MORES, supplied by the *Instituto Valenciano de Investigaciones Económicas* (IVIE) and the Spanish Ministry of Finance. This dataset provides regional accounting-type data on gross value added, labour compensation (including an imputation for self-employed), capital stock, total employment, wage-earners, and the unemployment rate.⁹

Full details on each variable, and its respective source, are given Table 2. Employment (n_i) is the log of total employees in region i ; k_i is the log of total regional capital stock; pr_i is regional productivity, measured as the log of total regional gross value added to the number of employees; and w_i is the log of the real wage. To capture the effects of aggregate product demand on labour demand, we include the ratio of regional consumption over regional GDP (cs_i) as a regional control variable. The labour demand equation also includes several nation-wide exogenous variables: *open* is the log of openness, defined as total exports plus total imports over GDP, while r is the real interest rate, defined as the difference between the nominal three-month interest rate minus the inflation rate, which in turn is defined as the one-year percentage change in the GDP deflator. The degree of openness measures the impact of increased trade flows on labour demand due to globalization, whereas the interest rate aims at proxying the real user cost of capital. In the wage-setting equation we also consider several nation-wide control variables, capturing wage-push determinants. These are social security benefits as a percent of GDP (b), the real minimum wage (mw), real oil prices (oil), and the unemployment rate (u) as a measure of labour market tightness.

We have tried several additional controls, such as financial wealth (measured as the ratio of the Madrid stock exchange index over productivity), direct and indirect taxes as a percent of GDP, or the real effective exchange rate. None of them were significant in our estimations. Furthermore, we explored an alternative to the use of the real interest rate, such as the change in real money balances. Results were similar to those reported below, which include the real exchange rate.

A final word of caution must be given. When empirically working with the labour

⁹The Spanish Labour Force Survey has undergone major methodological breaks during our period of analysis, which precludes the use of long time series. One of the advantages of the BD-MORES is that it provides homogeneous data on labour market variables for the 1980-2011 period.

share we should recall that the definition of the labour share is

$$ls = \frac{\text{wage bill}}{GDP} = \frac{\text{wage bill} / \text{employees}}{GDP / \text{employees}} = \frac{\text{avg. wage}}{pr} \quad (5)$$

In our case, the wage bill provided by the BD-MORES includes an imputation for the labour income of the self-employed, which makes the adjustment suggested by Gollin (2002) or Karanassou and Sala (2014) unnecessary. In addition, the standard measure of GDP is provided at market prices, which includes taxes on production and imports, but not subsidies. Since these taxes and subsidies are not regarded as components of generated income, they need to be excluded from the definition of GDP. This is the reason why we make use of the GDP at basic prices (also provided by the BD-MORES).

For the reasons given, the labour income share is computed according to equation (5). Subsequently, we retrieve the average wage as $W = \frac{ls*Y}{N}$, where Y is GDP and N is total employment (not in logs). Therefore, our measure of the labour share can be regarded as the wage-productivity gap, $w - pr$, and is fully consistent with the variables used in the estimation of the wage setting equation.

Table 2. Definitions of variables.

REGIONAL VARIABLES		NATIONAL VARIABLES	
ls_i	labour share ($= w_i - pr_i$)	$open$	openness (%)
n_i	total employment	r	real interest rate (%)
k_i	total capital stock	b	social security benefits (% of GDP)
w_i	average real total compensation	mw	real minimum wage
pr_i	labour productivity	oil	real oil prices
cs_i	private consumption (% of GDP)	u	unemployment rate (%)

Notes: all variables in logs unless otherwise indicated.

Sources: OECD ($open$, b , mw), IMF (oil), BD-REMS (r) and BD-MORES (rest of the variables).

3.2 Empirical methodology

Estimation of equations (3) and (4) involves several steps. We follow the ARDL approach to cointegration analysis (also known by the bounds testing approach) proposed by Pesaran (1997), Pesaran and Shin (1999) and Pesaran *et al.* (2001), which overcomes many of the problems of the traditional approach to the analysis and identification of long run relationships, especially those related to unit root tests (see inter alia Cochrane, 1991; or Perron and Ng, 1996). These authors show that tests for unit roots have low power in finite samples against the local alternative of a root close to, but below unity (Cochrane, 1991). In this context, the main advantage of relying on the ARDL approach as a testing

and estimation strategy is that it can be applied irrespective of whether the involved regressors are stationary or not, and therefore can avoid the pre-testing problems associated with the standard cointegration analysis.

The initial step consists in identifying the optimal lag structure of each model. Thus, for each group of Spanish regions, we construct a panel and run a first estimation of equations (3)-(4) by OLS. Then we use standard statistical information criteria (e.g., the Akaike Information Criterion, AIC) to determine the optimal lag length for each endogenous and exogenous variables. However, the presence of the lagged endogenous variable on the right hand side introduces an additional issue. As it is well known, the standard fixed effects estimator is consistent in dynamic panels with constant slopes as $T \rightarrow \infty$ for fixed N . However, the literature has shown that when T is small relative to N , the OLS may become inconsistent (for instance, when $N \rightarrow \infty$ for fixed T , it gives rise to biases as shown by Nickell, 1981). The demeaning process, which subtracts the mean value of the endogenous and each exogenous variable, creates a correlation between the regressor and the error term, which is not mitigated by increasing N , or by increasing the number of regressors. In this case, the standard procedure in the literature is to first difference the data to eliminate the fixed effects, and apply the General Method of Moments (GMM) estimator, as suggested by Arellano and Bond (1991) or Blundell and Bond (1998).

Consequently, we estimate the selected specification of the labour demand and wage setting equations by GMM to identify the short run and long run elasticities by group of regions. The former refers to the estimated coefficients from equations (3) and (4), whereas the latter refers to the long run solution of the models. This involves assuming that growing variables stabilise around a long run value, so that the long run solution of the dynamic models may be derived as:

$$n_i^* = \frac{v_i^N}{1 - \sum_j \alpha_j^N} + \frac{\sum_j \beta_j^N}{1 - \sum_j \alpha_j^N} l s_i^* + \frac{\sum_j \gamma_j^N}{1 - \sum_j \alpha_j^N} k_i^* + \frac{\sum_j \delta_j^N}{1 - \sum_j \alpha_j^N} \mathbf{x}_i^{N*} \quad (6)$$

and

$$w_i^* = \frac{v_i^W}{1 - \sum_j \alpha_j^W} + \frac{\sum_j \beta_j^W}{1 - \sum_j \alpha_j^W} p r_i^* + \frac{\sum_j \delta_j^W}{1 - \sum_j \alpha_j^W} \mathbf{x}_i^{W*} \quad (7)$$

where an asterisk denotes the long run value of the variable. These solutions measure the long run impact of a unit change in each exogenous variables, once the dynamics have been settled down.¹⁰

¹⁰See Karanassou *et al.* (2010) for a discussion about short run, long run and steady state solutions of

3.3 Estimates

Tables 3 and 4 summarise our results for the labour demand and wage-setting equations. Each table supplies the results for the High and Low income groups of regions, and the corresponding OLS and GMM estimates. In general, all variables are statistically significant at the standard confidence levels and show the expected signs.

The validity of the instruments in the GMM approach can be tested by the Sargan test, as well as the Arellano-Bond $M1$ and $M2$ correlation tests. Tables 3 and 4 show that the p -value for the Sargan tests does not allow rejection of the null that the over-identifying restrictions are valid. Moreover, the $M1$ and $M2$ statistics indicate that residuals in both equations, and for both groups of regions, show first-order but not second-order autocorrelation. Finally, both models provide an excellent fit for employment and the real wage in both groups of regions, as depicted by Figure 1.

Given this evidence, our models can be considered as appropriate representations of labour demand and wage setting conditions in the High and Low income regions.

The overall picture that emerges from these estimates, together with the long run solutions of the models (summarised in Tables 5 and 6), is the existence of a dual regional employment growth pattern.

Starting with the labour demand equation, we observe that the degree of persistence is rather large and similar in both groups: 0.84 and 0.91 in the high and low income groups, as obtained from the addition of the two lagged dependent variables. This is a crucial feature which, on one side, reflects some larger average adjustment costs in the low income regions; and, on the other, implies that differences in the long-run impact of the variables will mainly arise from the short-run coefficients on those variables –recall equations (6) and (7). Note, for example, that the addition of the labour share coefficients implies a much larger sensitivity of employment to this variable in the high income regions. This is then translated into the long-run elasticities of -0.94 and -0.49, respectively, as shown in Table 5. In turn, the level of the capital stock is only significant in the high income group, while the impact of investment (Δk) is significant in both groups, but with a larger sensitivity in the low income regions.

A dual pattern is thus identified. A dual pattern in which the high income group is more reactive to the widening of the wage-gap (the fall in the labour share), and the low income group is more sensitive to investment. This may be related to different technological levels at the regional level. It is likely that investment generates a greater income and employment effect on the low income regions due to larger economies of scale.¹¹

a dynamic equation.

¹¹Bande and Riveiro (2017) also find that investment generates larger GDP and employment effects in Spanish lagging regions.

Table 3. Labour demand equation.

	HIGH INCOME REGIONS				LOW INCOME REGIONS			
	<i>OLS</i>		<i>GMM</i>		<i>OLS</i>		<i>GMM</i>	
	coef.	p-val.	coef.	p-val.	coef.	p-val.	coef.	p-val.
n_{t-1}	1.336** (21.23)	0.00	1.298** (20.58)	0.00	1.294** (26.21)	0.00	1.216** (27.35)	0.00
n_{t-2}	-0.490** (-8.21)	0.00	-0.462** (-7.81)	0.00	-0.348** (-7.22)	0.00	-0.307** (-7.82)	0.00
ls_t	0.244** (4.56)	0.00	0.225** (4.20)	0.00	0.118** (2.66)	0.00	0.091** (2.86)	0.00
ls_{t-1}	-0.365** (-6.79)	0.00	-0.380** (-7.08)	0.00	-0.142** (-3.19)	0.00	-0.136** (-2.99)	0.00
k_t	0.066** (4.30)	0.00	0.070** (4.50)	0.00	0.002 (0.26)	0.78	0.021 (0.98)	0.32
Δk_t	1.104** (5.19)	0.00	1.085** (5.20)	0.00	1.182** (7.87)	0.00	1.315** (5.85)	0.00
Δk_{t-1}	-0.549* (-2.49)	0.01	-0.451* (-2.06)	0.04				
Δk_{t-2}					-0.871** (-6.16)	0.00	-0.804** (-4.82)	0.00
$open_t$	0.218** (4.81)	0.00	0.229** (5.16)	0.00				
$open_{t-1}$	-0.230** (-4.37)	0.00	-0.224** (-4.35)	0.00				
$\Delta open_t$					0.200** (5.58)	0.00	0.186** (6.87)	
r_t	-0.421** (-5.78)	0.00	-0.421** (-5.93)	0.00				
r_{t-1}	-0.155** (-2.57)	0.00	-0.162** (-2.73)	0.00	-0.320** (-6.88)	0.00	-0.321** (-6.64)	0.00
c_{t-1}	0.205** (4.04)	0.00	0.260** (4.87)	0.00	0.063** (1.52)	0.12	0.120** (2.46)	0.01
	<i>LL</i>	571.91	<i>Sargan</i>	187.05 (0.422)	<i>LL</i>	767.85	<i>Sargan</i>	229.08 (0.13)
	<i>AIC</i>	-5.265	<i>m1</i>	-7.51 (0.00)	<i>AIC</i>	-5.15	<i>m1</i>	-7.53 (0.00)
			<i>m2</i>	1.73 (0.10)			<i>m2</i>	0.994 (0.31)

Notes: OLS refers to the one-way fixed effects model. GMM refers to the Arellano-Bond first-difference generalised method of moments. Standard-errors in parentheses for estimated coefficients. *LL* is the maximised value of the log-likelihood function, *AIC* is the value of the Akaike Information Criteria. *Sargan* is the value of the Sargan test. *m1* and *m2* refer to the Arellano-Bond autocorrelation test for residuals. p-values in parenthesis for the *Sargan*, *m1* and *m2* tests. ** significant at 1%; * significant at 5%.

The remaining variables in the labour demand equations behave in line with theoretical priors. Real interest rates have the expected negative influence on employment, which is remarkably similar across groups of regions (note that both estimated long run

elasticities are around -3.5). In turn, consumption impacts positively on labour demand, reflecting significant aggregate product demand effects which, again, have similar long-run elasticities (1.58 and 1.33 for the high and low income regions). Finally, the degree of openness impacts positively on employment only in the high income group of regions (0.03), probably reflecting their better level of competitiveness.

Table 4. Wage setting equation.

	HIGH INCOME REGIONS				LOW INCOME REGIONS			
	OLS		GMM		OLS		GMM	
	coef.	p-val.	coef.	p-val.	coef.	p-val.	coef.	p-val.
w_{t-1}	0.752** (12.85)	0.00	0.753** (12.92)	0.00	0.687** (14.88)	0.00	0.601** (12.52)	0.00
w_{t-2}	0.128* (2.21)	0.02	0.132* (2.31)	0.02	0.145** (3.45)	0.00	0.144** (3.09)	0.00
pr_t	0.463** (7.58)	0.00	0.462** (7.62)	0.00	0.483** (8.11)	0.00	0.460** (7.87)	0.00
pr_{t-1}	-0.407** (-6.70)	0.00	-0.415** (-6.87)	0.00	-0.375** (-8.68)	0.00	-0.310** (-6.04)	0.00
b_t					2.132** (7.40)	0.00	2.366** (7.12)	0.00
b_{t-1}					-1.820** (-6.85)		-1.446** (-5.99)	0.00
Δb_t	1.41** (5.06)	0.00	1.423** (5.20)	0.00				
u_t					-0.067 (-1.38)	0.16	-0.175** (-2.47)	0.01
mw_t	0.094** (2.76)	0.00	0.094** (2.81)	0.00	0.067 (1.46)	0.14	0.022 (0.51)	0.60
oil_t	-0.011** (-3.41)	0.00	-0.011** (-3.52)	0.00				
oil_{t-1}					-0.018** (-5.19)	0.00	-0.017** (-5.30)	0.00
LL			546.19	Sargan	188.00 (0.40)	LL		
AIC			-5.068	$m1$	-8.315 (0.00)	AIC		
				$m2$	0.793 (0.42)	$m1$		
						243.59 (0.10)		
						$m2$		
						-8.561 (0.00)		
						$m2$		
						-0.435 (0.66)		

It just attains 0.41 in the high income regions and 0.58 in the low income ones. This is essentially the outcome of the larger short-run sensitivity in the latter (0.15 versus 0.05), and also of its mild smaller persistence (0.754 versus 0.885, arising from the addition of the two lags of the real wage).

The control variables have also a diverse impact across regions. Minimum wages appear as a much relevant wage pressure factor in the high income regions than in the low income ones. This is consistent with their larger sensitivity of employment to the wage gap. In contrast, it is not the level of Social Security benefits what impacts in the rich regions, but its change (hence the absence of a long-run elasticity with respect to the level). In the low income regions, on the contrary, the long-run elasticity of wages with respect to benefits is high (3.62). This larger pressure of benefits is probably reflecting greater labour market difficulties in these regions which, in turn, would be reflected in the wage response to the rate of unemployment. Wages react significantly to higher rates of unemployment in the low income regions, but not in the rich ones.

Table 5. Labour demand. Long run elasticities.

		<i>ls</i>	<i>k</i>	Δk	<i>open</i>	<i>r</i>	<i>cs</i>
HIGH INCOME REGIONS:	<i>OLS</i>	-0.779	0.432	3.598	-0.075	-3.738	1.330
	<i>GMM</i>	-0.943	0.425	3.855	0.030	-3.548	1.585
LOW INCOME REGIONS:	<i>OLS</i>	-0.434	0.051*	5.754	-	-5.948	1.180*
	<i>GMM</i>	-0.493	0.234*	5.626	-	-3.541	1.330

Notes: values computed from estimated coefficients in Table 3; * denotes non-significant level coefficient.

Table 6. Wage equation. Long run elasticities.

		<i>pr</i>	<i>mw</i>	<i>b</i>	<i>u</i>	<i>oil</i>
HIGH INCOME REGIONS:	<i>OLS</i>	0.469	0.799	-	-	-0.098
	<i>GMM</i>	0.412	0.826	-	-	-0.104
LOW INCOME REGIONS:	<i>OLS</i>	0.639	0.404*	1.861	-0.402*	-0.109
	<i>GMM</i>	0.587	0.087*	3.620	-0.690	-0.067

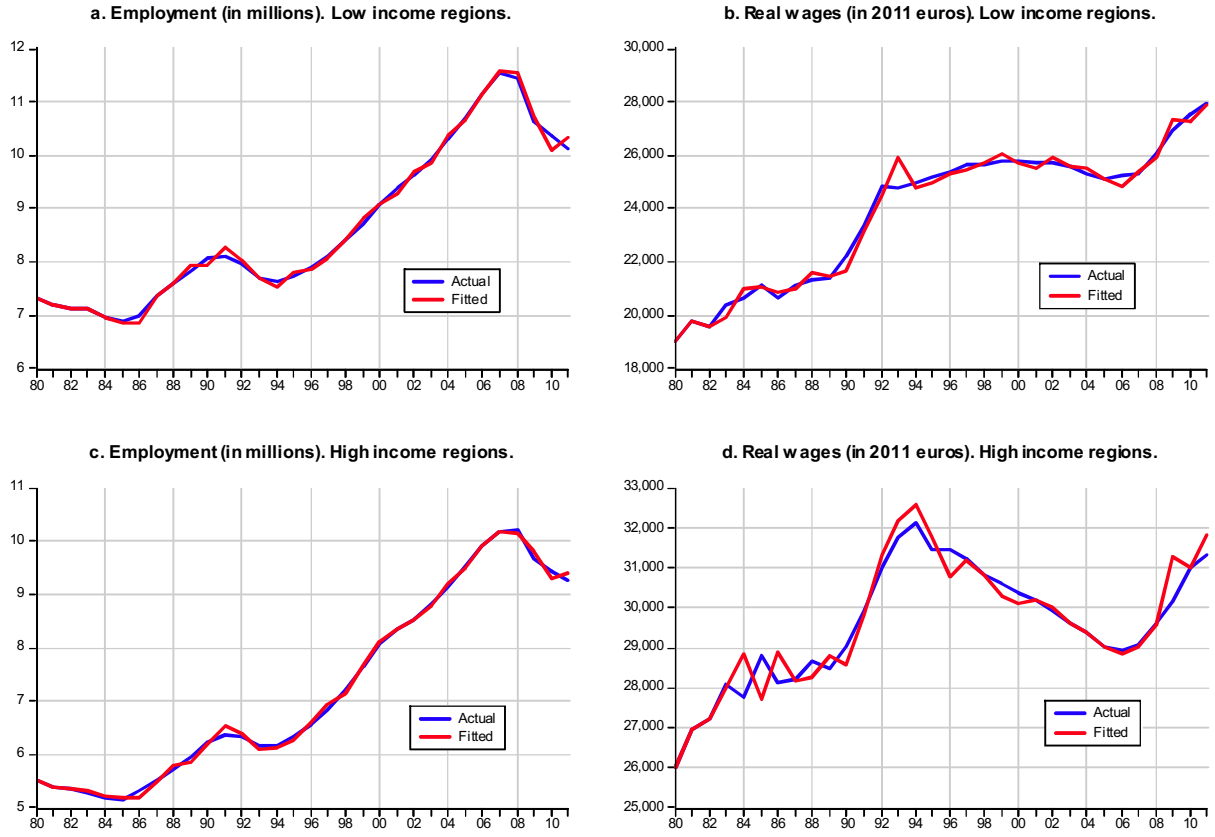
Notes: values computed from estimated coefficients in Table 4; * denotes non-significant level coefficient.

These results suggest the existence of a dual pattern in the Spanish regions, both in terms of employment growth and wage setting, inspite of sharing the same institutional setting. Moreover, our findings confirm and reinforce those in Bande *et al.* (2008, 2012).

Figure 1 shows the fitted values corresponding to the employment and real wage equations estimated for the low and high income regions. The fitted values track very closely the actual evolution of the explained variables, with mild discrepancies only in the 1980s for real wages in the high income regions. Note, however, that our simulations

comprise the 1990s and 2000s and are therefore based on a faithful representation of the facts.

Figure 1. Fitted values.



4 Learning from counterfactuals

Having found that the labour share and investment impact quite differently on the High and the Low income groups of regions, we may wonder what has been the contribution of these variables to the actual evolution of employment. To this end, and following the methodology proposed by Karanassou and Sala (2014), we now run a number of counterfactual simulations. These consist in fixing the value of one of these exogenous variables at a certain level, and solve dynamically the estimated labour demand or wage setting model, allowing the remaining exogenous variables to take their actual values. The corresponding simulated employment level may be interpreted as the level of employment that the group of regions would have showed had the exogenous variable under scrutiny remained constant. In addition, the difference between the two (actual and simulated trajectories) is the dynamic contribution of that exogenous variable to the trajectory of employment or wage setting.

4.1 Counterfactual scenarios

We focus on three significant periods: 1993-1999, 1999-2007, and 2007-2011. The first two cover a long span of job creation (starting in 1995) and coincide with significant institutional and regulatory changes such as the achievement of the Common Market, in the first case; the achievement of the European and Monetary Union in the second; and a variety of labour market and fiscal reforms that plagued those years. In turn, the third period covers the aftermath of the Global Financial Crisis (GFC) with the burst of the housing bubble and the collapse of the saving banks (which accounted for more than half of the financial sector in Spain in terms of credits, deposits and customers).

Figure 2. Counterfactuals.

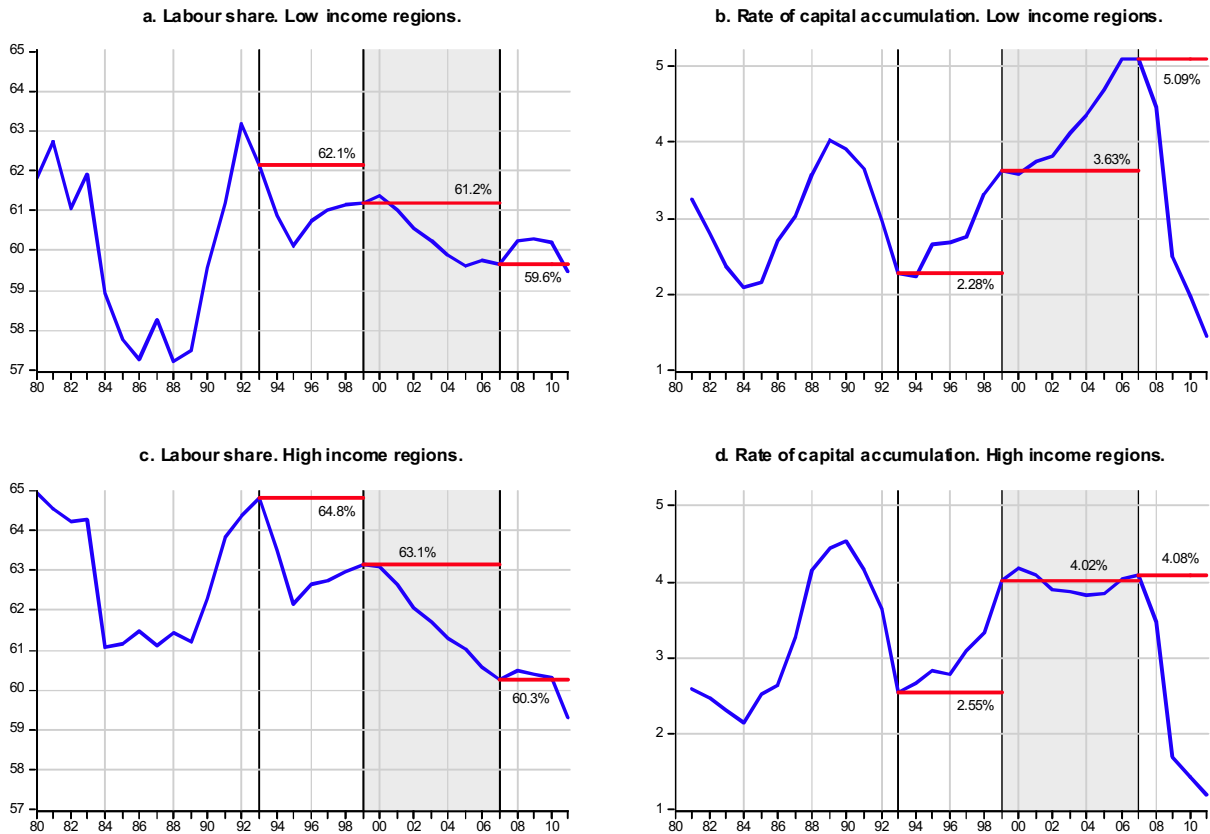


Figure 2 shows that 1993, 1999 and 2007 broadly coincide with inflection points in the time path of the labour share and the rate of capital accumulation. Regarding the labour share, 1992 and 1993 were the peak years resulting from the break in the income policy agreements that lasted until 1987, and the subsequent period of real wage growth (see Figures 1b and 1d).¹² Therefore, when fixing the values of the labour share in 1993

¹²After democracy (1977), and up to 1987, different agreements between representatives of firms, trade unions and the government were in place to break the wage-price spiral brought by the oil price shocks. The result was a slow progress in real wages in a context of intensive job destruction. Hence the fall in the labour share. When sustained growth was regained in the second half of the 1980s, income policy

we will be able to evaluate the impact of its fall during the three periods examined. On this account, note that the shock is larger in the high income regions (Figure 2c), where it falls by 1.7 percentage points (pp) between 1993 and 1999 (0.9 in the low income regions as shown in Figure 2a); 2.8 pp in 1999-2007 (1.6 in the low income regions); and 1 pp in 2007-2011 (0.1 in the low income regions) after increasing mildly, along with the collapse in output, during 2008-2010.

With respect to the rate of capital accumulation, there is a first, long span of accelerating growth rates, followed by a steep deceleration along with the burst of the housing bubble. In contrast to the changes in the labour shares, the low income regions display wider oscillations than the high income ones. In the first period, the rate of capital accumulation changes by a similar magnitude in both areas: by 1.35 percentage points in the low income regions (from 2.28% to 3.63%) and 1.47 in the high income ones (from 2.55% to 4.02%). However, in the second and third periods both the rise and fall of these rates are larger. There is an increase of 1.46 pp in 1999-2007, in contrast to the flat trajectory in the high income regions, while there is steep fall of 3.69 pp in 2007-2011 (from 5.09% to 1.50%), larger than the 2.88 pp deceleration recorded in the high income regions.

4.2 Employment simulations

Figures 3 and 4 show the results of the dynamic simulations for the employment trajectories corresponding, respectively, to the low and high income regions.

Employment grew by 1.0 million between 1993 and 1999 in the low income regions, and 1.5 millions in the high income regions. The labour share and capital accumulation jointly account for between a third and 40% of this evolution, with significant differences across areas.

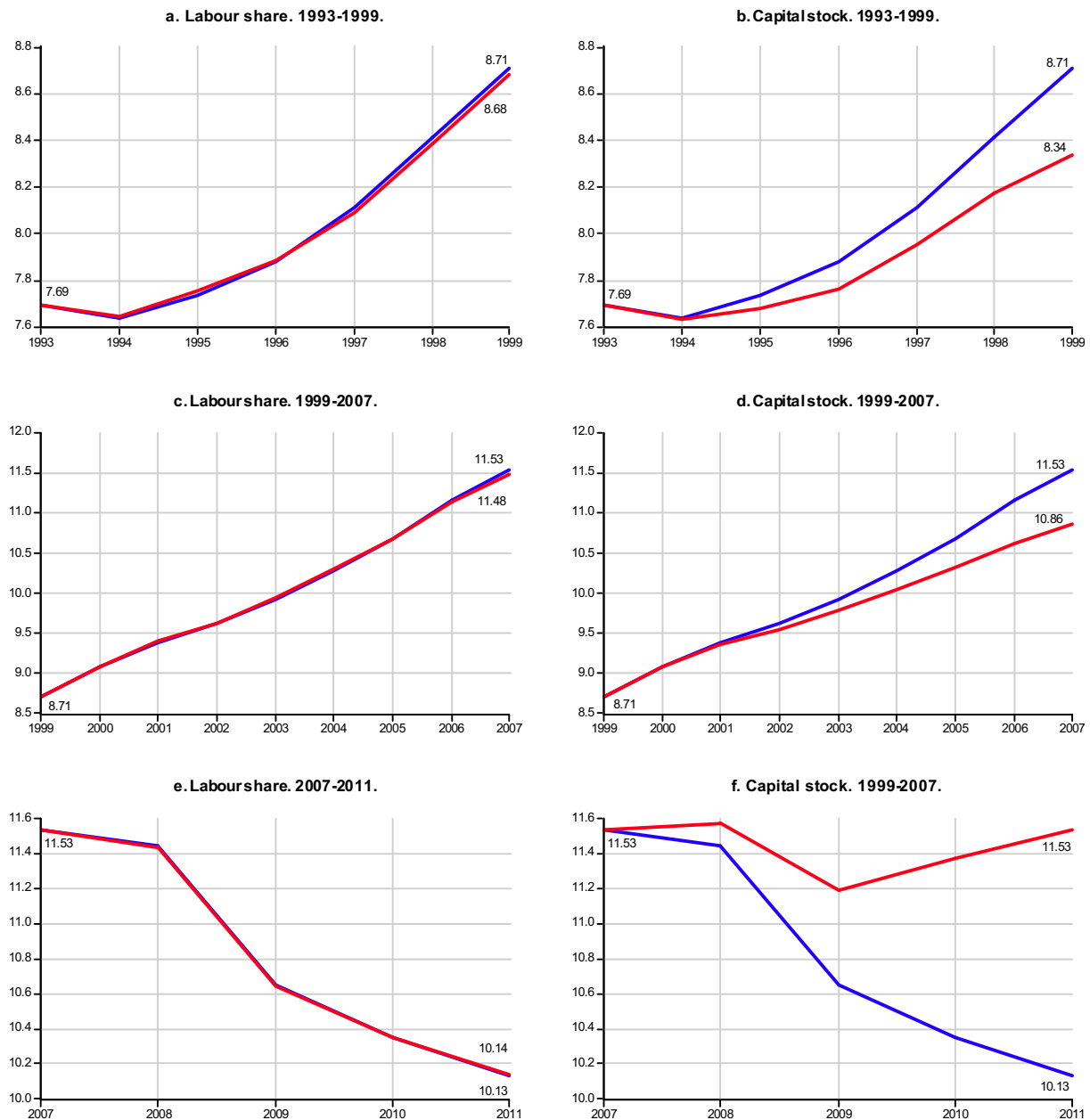
In the low income regions, the incidence of the labour share is negligible, as shown by Figure 3a, and only accounts for 3% of the progress in employment (8.71-8.68 millions, or 30,000 jobs, out of 1 million increase). In contrast, the acceleration in the rate of capital accumulation explains 37% of the job creation (8.71-8.34 millions, or close to 370,000 jobs). These contributions are much more balanced in the high income regions, where the labour share explains 18.4% of the jobs created in the high income regions (around 270,000), while capital accumulation accounts for another 15.5% (close to 230,000 jobs). In total, half million jobs out of 1,5 millions in this area.

The contribution of the labour share in the low income regions remains negligible in 1999-2007 (Figure 3b). In turn, the acceleration in the rate of capital accumulation accounts for 23.9% of more than 2.8 million jobs created in this area, which amounts to 670,000 jobs (11.53 - 10.86 millions, as shown in Figure 3d). This outcome is somewhat

agreements came to an end, and the massive general strike in 1988 gave rise to a new period of real wage growth.

reversed in the high income regions, where the fall in the labour share accounts for 360,000 jobs (10.18-9.82) out of 2,5 millions (14.4%), while there is virtually no contribution from changes in the capital stock. This is the outcome of the stability in the rate of capital accumulation around 4% (Figure 2d). We are not claiming, therefore, that the boom in investment was innocuous to job creation; the exercise just highlights the fact that low income regions did benefit more from the continued acceleration in the rate of capital accumulation, in contrast to the high income regions, where the flat evolution of that rate did not contribute by more than it was doing in 1999.

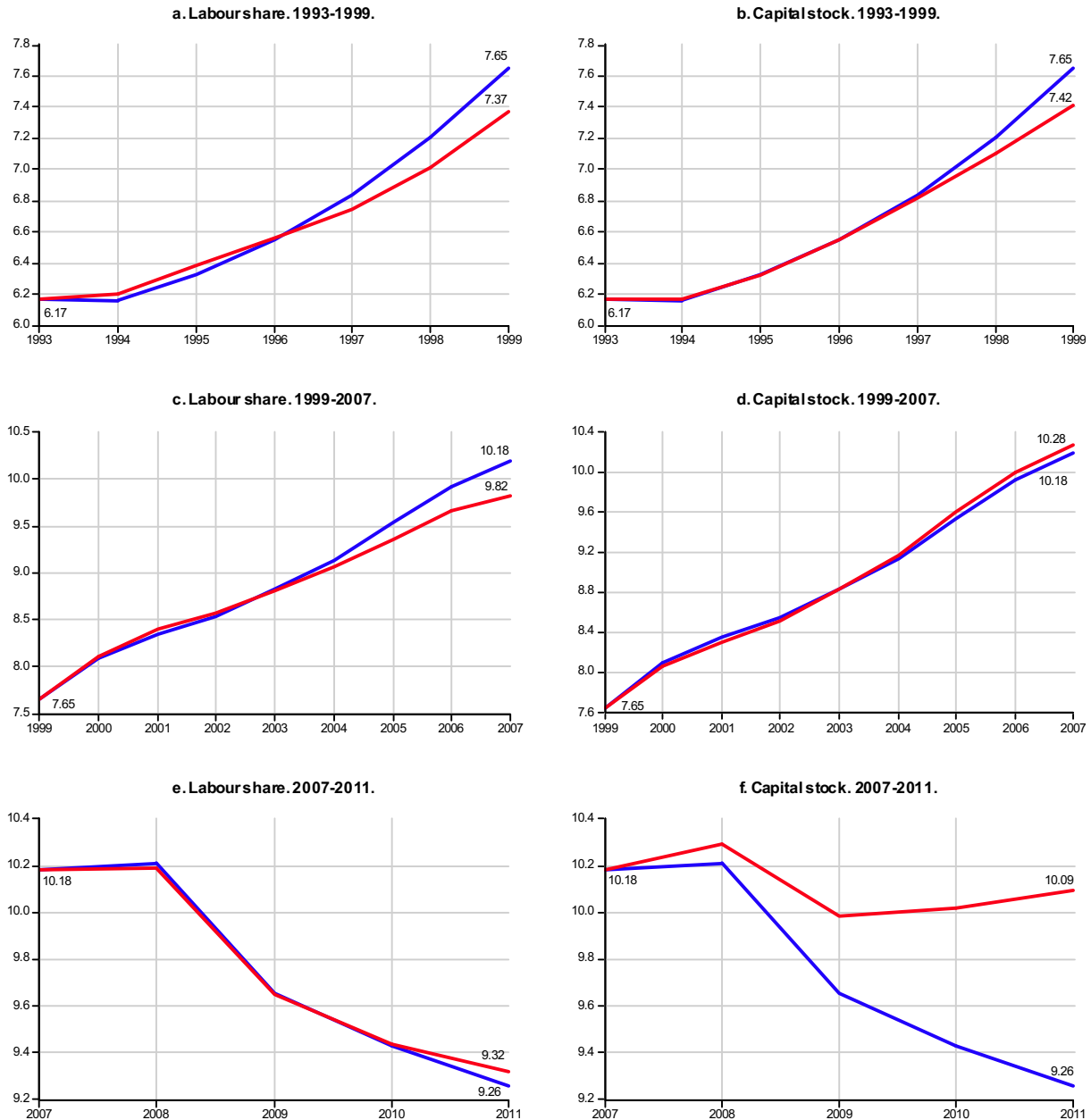
Figure 3. Dynamic simulations for employment. Low income regions.



In contrast to the dissimilar experiences of 1993-1999 and 1999-2007, the third period reveals a common behaviour. The evolution of the labour share is non-influential even in

the high income regions (recall the little magnitude of the shock, as depicted in Figures 2b and 2d), while fading investment and the corresponding steep fall in the rate of capital accumulation is able, on its own, to fully explain the employment trajectories (Figures 3f and 4f).

Figure 4. Dynamic simulations for employment. High income regions.



More precisely, it explains 100% of the fall in the low income regions, since employment would have remained exactly at 11.53 millions in the absence of the collapse in investment. In contrast, actual employment went down by 1.4 millions to 10.13. In the high income regions, it explains 90.2% of the total fall in employment; that is, more than 800,000 jobs out of an actual loss of 920,000 (10.18-9.26 millions).

Overall, the loss in Spain of 2.2 million jobs in those years (1.4+0.8) can be fully ascribed to the fall in investment. We know that the big collapse in investment took place immediately afterwards the GFC. Therefore, the years up to 2011 depict a period in which the rate of capital accumulation exerted utmost influence on economic activity and job creation. Subsequently, austerity policies were further emphasized with the extensive 2012 labour market reform. This may have had significant influence on the recent labour market performance and requires further analysis. Meanwhile, our estimated models of wage setting can be used to highlight the role played by labour productivity on the wage trajectories.

4.3 Real wage simulations and employment effects

The most crucial finding of our wage setting estimates is the absence of a one-to-one long run relationship between wages and productivity. If this relationship was to hold, the labour share would stay constant (in the long run), and in a Cobb-Douglas setting no jobs would be gained from improving labour productivity. What would be the consequence for wages of such one-to-one relationship? To respond to this question, we force the current coefficient on productivity to take a value such that this relationship holds (while maintaining the estimated coefficients on lagged productivity). Hence, we use the following simulation elasticities instead of the estimated ones:¹³

Estimated elasticities:		Simulation elasticities:	
$\frac{\beta_1^W + \beta_2^W}{1 - \alpha_1^W - \alpha_2^W}$ [notation from equation (4)]		$\frac{\beta_1^{W'} + \beta_2^{W'}}{1 - \alpha_1^W - \alpha_2^W}$ so that $\varepsilon_{w-pr}^{LR'} = 1$	
HI:	$\varepsilon_{w-pr}^{LR} = \frac{\hat{\beta}_1^W + \hat{\beta}_2^W}{1 - \hat{\alpha}_1^W - \hat{\alpha}_2^W} = \frac{0.46 - 0.41}{1 - (0.75 + 0.13)} = \frac{0.05}{0.12} = 0.42$	\rightarrow	$\varepsilon_{w-pr}^{LR'} = \frac{0.53 - 0.41}{1 - (0.75 + 0.13)} = \frac{0.12}{0.12} = 1.0$
LI:	$\varepsilon_{w-pr}^{LR} = \frac{\hat{\beta}_1^W + \hat{\beta}_2^W}{1 - \hat{\alpha}_1^W - \hat{\alpha}_2^W} = \frac{0.46 - 0.31}{1 - (0.60 + 0.14)} = \frac{0.15}{0.26} = 0.58$	\rightarrow	$\varepsilon_{w-pr}^{LR'} = \frac{0.57 - 0.31}{1 - (0.60 + 0.14)} = \frac{0.26}{0.26} = 1.0$

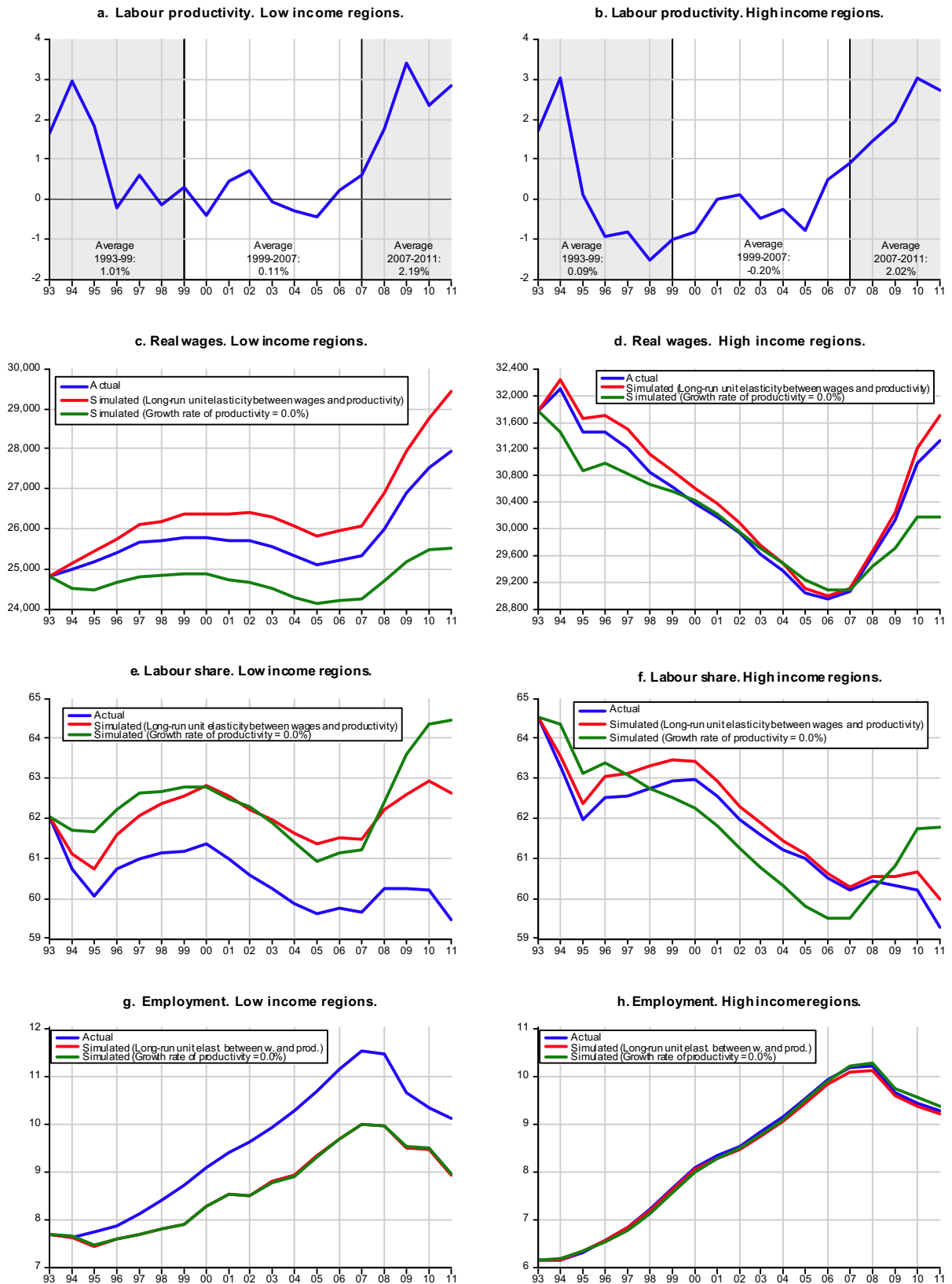
Then, we explore how real wages would have behaved in case of zero productivity growth. This provides us with two extreme scenarios, one in which wages are fully responsive to actual productivity gains, and one in which there are no productivity gains and thus the elasticity of wages with respect to productivity becomes irrelevant. These wage trajectories have strong consequences for the labour share, and can be used to forecast the evolution of employment to see how it would react to policies that close the wage gap (and push towards a one-to-one relationship), or leave wages fully independent from productivity changes.

Figure 5 presents information on the evolution of labour productivity together with the results of the dynamic simulations for the real wage, labour share, and employment

¹³ ε_{w-pr}^{LR} denotes the long-run elasticity of wages with respect to productivity; HI denotes the High income regions; LI denotes the Low income regions.

trajectories emerging from these two simulated scenarios.

Figure 5. Dynamic simulations for real wages.



In Figures 5a-5b, we show that labour productivity has followed a similar path in

both the low and high income regions. There is, however, a difference in the magnitudes recorded in the first period, 1993-1999, in which productivity grew at 1.0% in the former and 0.07% in the later. This stark contrast in the first part of the sample causes the average growth rate of productivity in the low income regions to double the one in the high regions: 0.95% versus 0.46% in 1993-2011.

Figures 5c and 5d show how real wages would have evolved had the growth rate of productivity been 0.0% since 1993. In the low income regions, average real wages would have increased by 3.9% (from close to 24,800 to more than 25,500 euros) instead of the actual 12.8% (which took them to 28,000 euros). In the high income regions, wages would have decreased by 4.3% (from 31,778 to 30,399 euros) instead of falling by just 1.4% (until 31,328 euros). This information illustrates, with specific values, what we learnt from the estimated elasticities: since not all productivity gains end up translated into real wages, the wage gap widens and the labour share falls.

If we take the actual evolution of productivity, but assume a one-to-one long run relationship between wages and productivity, then wages would reflect to a greater extent all productivity gains which, as we have seen, have been very low but still positive since the 1990s. In that case, wages would have increased by 18.8% (to 29,445 euros) in the low income regions, while they would have fallen in the high income regions by 8.7% until 2006, to virtually regain their initial values by the end of the period. This difference is partly due to their different productivity dynamics, with mild positive rates in the low income regions and negative rates in the high income regions in 1995-2007.

This analysis yields one significant outcome: the low income regions, which are the ones where the wage elasticity with respect to productivity is larger and the growth rate of productivity has been higher, would be the most sensitive ones to policies that: (i) increase labour productivity; and (ii) increase the sensitivity of wages with respect to productivity. The first type of policy could be a growth-type policy aiming at improving technology (the A in a production function), and thus labour and capital efficiency. The second type of policy could be a change in legislation promoting wage agreements in which wage progress becomes more connected to the workers' performance in terms of productivity (for example, the 2012 labour market reform in Spain).¹⁴ What would be the consequences for employment?

Figures 5e and 5f show the consequences of our wage scenarios for the labour share, while Figures 5g and 5h show them for employment. In the scenario of flat productivity growth, the wage gap narrows and the labour share evolves systematically above its actual trajectory: between 61% and 63% until 2008, then reaching 64.5% in 2011. This implies 5 pp above its actual value. In contrast, the zero productivity growth causes the labour share

¹⁴See the OECD (2014) assessment of this reform, where the greater priority given to collective bargaining agreements at the firm level is noted as a key measure, and is highlighted as crucial to enhance the internal flexibility of firms.

to display a steeper falling trend during most of the sample period (it ends up above its actual level due to the steep rise that would have occurred in the absence of productivity growth during the crisis, which is similar to the one in the low income regions).

The narrowing of the wage gap in the low income regions, and its widening (in most of the sample period) in the high income regions, has distinct effects on employment. In the first case (depicted in Figure 5g), it is much lower than actual employment (by more than 1 million in 2011), as a consequence of the higher unit labour costs due to stagnant productivity. In the second case (Figure 5e), the evolution of simulated employment closely tracks actual employment, and no significant effects can be perceived. In the high income regions, not only the elasticity of wages with respect to productivity is lower (as reported in Table 6), but also the average productivity growth rate (Figure 5b), hence rendering causing this area to be much less sensitive in terms of employment.

These simulations show that policies trying to foster employment through a low productivity growth strategy –i.e., based on a low value added economic model– are deemed to fail. Their effects would tend close to neutral in developed regions, while employment would be reduced in more lagging regions. Thus, not only employment growth would be limited at best in some areas, but regional disparities would markedly increase.

The second scenario in which productivity gains end up fully translated into wages is equally eloquent. Now productivity takes its actual values (it grows), while wages reflect them more intensively (in the short-run) and fully (in the long-run). As a consequence, the simulated labour share is higher than the actual one in the low income regions (Figure 5e). It is remarkable that the resulting path is very similar to the one arising from the previous scenario but, given this similarity, it should come as no surprise to find a very similar employment response (Figure 5g).

In contrast, in the high income regions actual productivity gains have been, on average, about half those in the low income regions, with much of the difference occurring in the early 1990s (Figures 5a and 5b). This explains why, even though simulated wages are also higher, they do not detach much from their actual values (Figure 5d). This is reflected in the trajectory of the labour share, which is also above the actual one (reflecting the enhanced translation of productivity gains into wages), but not much (Figure 5f). Although for different reasons (before we had $\Delta pr = 0$ and $\varepsilon_{w-pr}^{LR} < 1$), this same outcome in terms of the labour share causes similar negligible employment effects.

This second scenario ($\Delta pr > 0$ and $\varepsilon_{w-pr}^{LR} = 1$), therefore, suggests that policies aimed at enhancing the link between wage and productivity growth (i.e., increasing wage responsiveness to productivity gains) may not have the desired effects, since they would cut employment growth in low income regions, while generating a neutral effects in high income regions. Focusing exclusively on the response of wages to productivity would widen regional disparities.

4.4 Discussion

The design of policies to foster employment is not an easy task, especially in the presence of marked regional disparities in labour market performance and a common set of labour market institutions. Our analysis sheds lights on the two main strategies that may be followed.

First, the implementation of growth-enhancing policies focusing on capital accumulation so that high productivity growth can be achieved. Under this approach, employment would be the natural outcome of the combined effect of an aggregate demand boost (due to greater investment) and aggregate supply developments from embodied technological change. Second, cost-control strategies which essentially try to keep wages in line with productivity growth, or gain competitiveness by allowing wages to fall short of productivity growth, i.e., by increasing the wage gap.

These alternatives are not mutually exclusive, and may be combined. In case of weak productivity growth, however, pursuing job creation by increasing the wage gap may require wages to be reduced and cause unavoidable effects on income distribution (Karanassou and Sala, 2014). Therefore, our results call for more region-specific policies and discard standard labour market reforms as a unique tool to manage the unemployment rate problem. Further, to the extent that investment serves both at fostering capital accumulation and labour productivity (which, in turn, reduces the ULC), regionally-targeted soft credit lines and capital taxes could be helpful in breaking regional sluggishness.

Achieving such policy goals, however, is certainly more difficult today than before for a twofold reason. On the one hand, the restructuring of the Spanish banking system after the GFC has swept regional saving banks, which use to have more regionally-oriented credit strategies. On the other, capital taxes are fully ruled out from the political debate at the national level.

What is left, then? Looking at the recent policies applied in Spain to tackle the high and persistent unemployment problem under the lens of our results does not allow to be optimistic as regards the future evolution of regional employment disparities, let alone aggregate employment. First, growth-enhancing policies have been essentially neglected, since job creation in Spain during the expansionary years before the GFC, and its aftermath (2014-2016), has been based on low value added activities such as building construction, tourism, wholesale and retail. The focus on labour-intensive industries has resulted in weak productivity growth and has prevented wages to progress, as implicitly signalled by our simulations. Wage growth recorded a yearly average of 0.64%, which would have been 0.15% in the absence of productivity growth in the low income regions (Figure 5c). In the high income regions, these rates are, respectively, 0.06% and -0.14% (Figure 5d).

Second, in a context of endemically slow productivity growth, labour legislation has

exerted systematic downward pressure on wages to secure job creation via the widening of the wage gap. To exploit this mechanism, the Spanish legislation has tended to foster the atomization of job categories. This has been done by allowing a large variety of contracts and job situations, in which only the compensation of the so-called insiders copes with progress in productivity, in contrast with a vast segment of temporary and outsourced workers liable to a multiplicity of (task-targeted) contracts. Even the few temporary work agencies that dominate the market use their monopsonistic power and enforce flat tariffs not specifically connected to the workers' performance.

In this context, changes in the legislation driving collective agreements to set wages much more in connection to progress in productivity (as expected from the 2012 reform in Spain), could even harm the situation unless significant reforms are undertaken to reduce workers' discrimination issuing from the type of contract. The reason is straightforward: in the event of a wide distribution of productivities across branches of activity and size of firms (Spain is an economy essentially made of SMEs), firm-level agreements connected to specific productivity performances in a context, in addition, of structural difficulties to achieve productivity gains, will surely deliver growing wage dispersion and income inequality. The recent promotion of firm-level agreements does not even guarantee that wages will trail productivity, thus precluding significant employment effects.

5 Conclusions

Labour market reforms are recursively invoked from academics and global institutions to solve the poor performance of the (Spanish) labour market. Policy makers in Spain have responded to this call by introducing systematic legislative changes in last decades (in 1984, 1994, 1997, 2001, 2002, 2010, 2011, and 2012) so as to fight the so-called unfriendly labour market institutions and enhance flexibility. As a consequence, the Spanish labour market displays worrying symptoms of a multiple personality disorder. On one side, it is highly volatile and very sensitive to business cycles; on the other side, regional asymmetries are extremely persistent.

In such disturbing context, our main claim is that a more targeted approach is needed. Targeted, to be precise, in a twofold dimension.

First, in regional terms. Although it is true that regions react differently to common policies due to compositional differences, such compositional differences need to be carefully identified so as to ensure the most efficient outcome of new, and increasingly targeted, economic policies.

Second, in orientation. While labour market performance may be more sensitive to legislation changes in some specific regions, other may require other type of measures more focused on growth stimulous. Examples of such measures are industrial policies

that care on R&D&I activities; policies that boost investment (via soft-credit lines or the promotion of joint ventures); and regionally-targeted tax incentives.

Our analysis has shown that the performance of the labour market in the Spanish high income regions is more reactive to capital accumulation, while in the low income regions is more sensitive to low costs (low relative to productivity). This is a useful piece of evidence, that of course will need to be further contrasted, but hopefully will contribute to change the perspective under which we tend to approach the labour market problem.

To conclude, it should be made explicit that our recommendations aim only at identifying the best mechanism to break regional (un)employment persistence in a labour market characterised by its volatility at the aggregate level. The design of policies aiming at regional convergence is far beyond the scope of our analysis, although coordination between the two would certainly be a plus.

References

- [1] Adam, A. and T. Moutos (2014): “Industry-Level Labour Demand Elasticities Across the Eurozone: Will There Be Any Gain After the Pain of Internal Devaluation?”, *CESifo Working Paper Series 4858*, CESifo Group Munich.
- [2] Antràs, P. (2004): “Is the U.S. Aggregate Production Function Cobb-Douglas? New Estimates of the Elasticity of Substitution”, *The B.E. Journal of Macroeconomics*, De Gruyter, vol. 4(1): 1-36, April.
- [3] Arellano, M., and S. Bond (1991): “Some Test of Specification for Panel Data: Monte Carlo Evidence and an Application to Employment Equations”, *Review of Economics Studies*, 58: 277-297.
- [4] Ball and Mankiw (2002): “The NAIRU in Theory and Practice”, *Journal of Economic Perspectives*, vol. 16(4): 115-136.
- [5] Bande, R., M. Fernández, and V. Montuenga (2008): “Regional unemployment in Spain: Disparities, business cycle and wage setting”, *Labour Economics*, vol. 15(5): 885-914.
- [6] Bande, R. and M. Karanassou (2009): “Labour market flexibility and regional unemployment rate dynamics: Spain 1980-1995”, *Papers in Regional Science*, vol. 88(1): 181-207.
- [7] Bande, R. and M. Karanassou (2014): “Spanish Regional Unemployment Revisited: The Role of Capital Accumulation”, *Regional Studies*, vol. 48(11): 1863-1883.
- [8] Bande, R., and D. Riveiro (2017): “Cohesion and Monetary Policies in the EU: Two Counteracting Forces? Evidence from the Spanish Experience”, IDEGA, *Working Paper*.
- [9] Blanchard and Katz (1999): “Wage Dynamics: Reconciling Theory and Evidence”, *AEA Papers and Proceedings* 89(2): 69-74.
- [10] Blundell and Bond (1998): “Initial Conditions and Moment Restrictions in Dynamic Panel Data Models”, *Journal of Econometrics* 87(1): 115-143.
- [11] Cochrane, J. (1991): “A Critique of the Application of Unit Root Tests”, *Journal of Economic Dynamics and Control*, vol. 15(2): 275-284.

- [12] Gollin, D. (2002): “Getting Income Shares Right”, *Journal of Political Economy*, vol 110(2): 458-474.
- [13] Hamermesh, D.S. (1993): *Labor Demand*, Princeton University Press, Princeton.
- [14] Hatton, T.J. (2007): “Can Productivity Growth Explain the NAIRU? Long-Run Evidence from Britain, 1871-1999”, *Economica*, 74: 475-491.
- [15] Hijzen, A. and P. Swaim (2010): “Offshoring, labour market institutions and the elasticity of labour demand”, *European Economic Review*, 54 8): 1016-1034.
- [16] Karanassou, M. and H. Sala (2014): “The Role of the Wage-Productivity gap in Economic Activity”, *International Review of Applied Economics*, vol 28(4): 436-459.
- [17] Karanassou M., Sala H., and D.J. Snower (2010): “Phillips Curves and Unemployment Dynamics: A Critique and a Holistic Perspective”, *Journal of Economic Surveys*, 24, 1-51.
- [18] Karabarbounis, L. and Neiman, B. (2014): “The global decline of the labor share”, *The Quarterly Journal of Economics*, 129(1): 61–103.
- [19] Krishna, P., Mitra, D. and Chinoy, S. (2001): “Trade liberalization and labor demand elasticities: evidence from Turkey”, *Journal of International Economics*, 55: 391-401.
- [20] Krishna, P., Shin, J. (2012): “Import protection, exports and labor-demand elasticities: Evidence from Korea”, *International Review of Economics and Finance*, 53: 91-109.
- [21] Lindbeck, A. and D.J. Snower (1994): “How are Product Demand Changes Transmitted to the Labor Market?”, *The Economic Journal*, vol. 104 (423): 386-398.
- [22] Lindbeck, A., and D.J. Snower (2001): “Insiders versus Outsiders”, *Journal of Economic Perspectives*, 15 (1): 165-188.
- [23] Nickell, S. (1981): “Biases in dynamic models with fixed effects”, *Econometrica*, 49(6), pp. 1417-26.
- [24] OECD (2014): *The 2012 Labour Market Reform in Spain: A Preliminary Assessment*, OECD Publishing, Paris. DOI: <http://dx.doi.org/10.1787/9789264213586-en>.
- [25] Perron and Ng (1996): “Useful Modifications to Some Unit Root Tests with Dependent Errors and their Local Asymptotic Properties”, *The Review of Economic Studies*, vol 6(3): 435-463.
- [26] Pesaran, H. (1997): “The Role of Economic Theory in Modelling the Long Run”, *Economic Journal*, vol. 107: 178-191.
- [27] Pesaran, M.H. and Y. Shin (1999): “An Autoregressive Distributed-Lag Modelling Approach to Cointegration Analysis” in Strom, S. (ed.), *Econometrics and Economic Theory in the Twentieth Century: The Ragnar Frisch Centennial Symposium*, Cambridge University Press: 371-413.
- [28] Pesaran, M.H., Shin, Y. and Smith, R.J. (2001): “Bounds testing approaches to the analysis of level relationships”, *Journal of Applied Econometrics*, 16: 289-326.
- [29] Rodrik, D. (1997): *Has Globalization Gone Too Far?*, Institute for International Economics, Washington, DC.
- [30] Shapiro, C. and J. Stiglitz (1984): “Equilibrium Unemployment as a Worker Discipline Device”, *The American Economic Review*, vol. 74(3): 433-444.
- [31] Slaughter, M. (2001): “International trade and labor-demand elasticities”, *Journal of International Economics*, 54: 27-56.

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