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The Political Economy of European Asylum Policies

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The Political Economy of European Asylum Policies

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The Political Economy of European Asylum Policies

Abstract

Despite the recognition that asylum policies are partly determined by political economy factors in the destination country, there is little empirical evidence on the precise linkages between those political factors and asylum policies. We shed light on this issue by examining the impact of elections and parties on first-time asylum applications. Our evidence is based on a large bilateral panel data set comprising 12 European destination countries and their 51 most relevant origin countries during the time period 2002 to 2014. Our findings suggest that the number of asylum applicants under left- and right-wing parties converges before elections and differs thereafter. This result is robust to several different specifications and suggests that both left- and right-wing cabinets choose moderate policies before the election and less moderate policies after the election.

JEL Code: H11, D72, F22.

Keywords: Electoral cycles, migration policies.

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1 Introduction

Since March 2016, the European Union has signed an agreement with Turkey to close off the migration route across the Aegean, one of the most popular routes for asylum-seekers from Asia. This has been done despite protests from many human rights groups that think such an agreement violates the international Convention on Refugees. In a tribune entitled “Keep them away”, *The Economist* (2017) describes policies intended by many EU countries as policies “where asylum-seekers would need to apply from abroad rather than coming to Europe.” This tension between the provision of asylum for rightfully entitled individuals and groups, and the pursuit of other political and economic goals has been acknowledged by the economic literature that compares the provision of asylum as a public good, where costs are privately assumed by the host countries (Moraga and Rapoport, 2014). With this background in mind, Dustmann et al. (2016) note that “the different exposures to refugee inflows and the lack of any effective European-level mechanism to ‘spread the burden’ of hosting refugee populations, led many countries to implement procedures aimed at reducing inflows into their territories.”

Despite the recognition that asylum policies are partly determined by political economy factors in the destination country, there is little empirical evidence on the precise linkages between those political factors and asylum policies. We shed light on this issue by examining the impact of elections and parties on first-time asylum applications. Our evidence is based on a large bilateral panel data set from Eurostat reporting quarterly origin-specific first-time asylum applications for 12 European destination countries and their 51 most relevant origin countries during the time period 2002 to 2014. We combine this dataset with the ParlGov dataset that comprises information on European national elections (date and outcomes) and party positions (Döring and Manow, 2016).

Our findings suggest that asylum policies are affected by the electoral cycle and the identity of incumbent parties. More precisely, we establish two main results: (i) before an election, the inflow of refugees is very similar across left and right cabinets; (ii) in the quarters following an election, the inflow of refugees diverges substantially, with significantly less asylum applicants under a right-wing cabinet. These patterns are robust to several different specifications and suggest that both left- and right-wing cabinets choose moderate policies before the election and less moderate policies after the election.

Our paper is linked to the literature on the determinants of refugee inflows.¹ In line with existing findings, in particular Hatton (2016), we confirm the importance of certain push factors in the origin country (e.g., political terror and civil war) as well as certain pull factors in the destination country (e.g., its labor market situation). Most importantly, we demonstrate the empirical importance of elections in the destination countries for asylum policies, in particular their relevance for the accessibility of these destination

¹See, e.g., (Czaika, 2009; Gudbrandsen, 2010; Hatton, 2009, 2016; Hatton and Moloney, 2015; Holzer et al., 2000; Moore and Shellman, 2007; Neumayer, 2004, 2005; Toshkov, 2014)

countries. Our political economy measures go beyond the existing literature's focus on the legal determinants of asylum flows and capture how asylum regulations are implemented de facto under changing governments² as well as during and off election periods. In doing so, our analysis relates to the literature on political budget cycles following Nordhaus (1975) which argues that incumbent politicians have strong incentives to distort public policies in order to increase approval rates whenever elections are pending. Our analysis can also be regarded as a test of whether parties converge to the interests of the median voter (Downs (1957)) or implement the policies they favor on ideological grounds (Hibbs (1977) and Alesina (1987))³.

The remainder of the paper proceeds as follows: Section 2 introduces the data used for our analysis. Section 3 presents the econometric framework, and section 4 discusses our main results. Then, section 5 provides several robustness checks. Finally, section 6 concludes.

2 Data

We use quarterly data on the number of origin-specific first-time asylum applications from a large bilateral panel data set from Eurostat. For our empirical analysis we select destination countries based on both data availability and size of the countries in terms of total first-time asylum applications in the time period 2002 to 2014. More precisely, our main specification comprises all countries which have data on origin-specific first-time asylum applications in at least 44 out of the 52 quarters under study and which report in total more than 30000 first-time asylum applications between 2002 and 2014.⁴ The list of destination countries and the total number of first-time asylum applications in these countries is presented in Table 1.⁵

[Table 1 about here.]

We further select the top 51 origin countries which together account for more than 90% of all first-time applications in the 12 destination countries between 2002 and 2014. As evident from Table 2 the top 10 origin countries already account for more than 45% of the total first-time applications.⁶

[Table 2 about here.]

²In a related study, Krause and Potrafke (2017) find evidence that the number of deportations of foreigners that are obliged to leave the country depends on the ideology of the government.

³For a detailed survey on partisan politics see Potrafke (2017).

⁴If countries have missing information in 2008 or 2009 we impute this data from data on the origin-specific total applications in the respective quarters. This information is only available from 2008 onwards and, therefore, it is not possible to impute missing information for the years before 2008. For the exact calculation see the data section in the online appendix.

⁵We exclude Cyprus due to several irregular cabinet changes.

⁶A full list of all 51 origin countries is provided in the online appendix.

Following Hatton (2016), we drop country pairs with very few applications in order to avoid cases with zero applications in many quarters. Thus, we keep only country pairs with at least two first-time asylum applications per quarter on average, which leaves us with 480 out of 612 possible origin destination combinations.

We combine this dataset with the ParlGov dataset that comprises information on European national elections (date and outcomes) and party positions (Döring and Manow, 2016). Figure 1 illustrates the time-line of elections and cabinet changes in the destination countries.⁷

[Figure 1 about here.]

The position of the government on a left-right scale is derived by weighting the left-right position of the parties in government against the ratio of the parties' seats to the cabinet's total seats in parliament. This procedure is illustrated by the following simple numerical example: Assume the government of country X is formed by a coalition of two parties, A and B, with A having 60 and B having 80 seats out of the 200 total seats in parliament. Assume further that A scores 4 and B scores 5 on the left-right score, where 0 indicates extreme left and 10 indicates extreme right. The left-right position of the cabinet is, then, calculated as follows:

$$LR\ score\ cabinet = LR\ score\ A * \frac{seats\ A}{seats\ A + seats\ B} + LR\ score\ B * \frac{seats\ B}{seats\ A + seats\ B}$$

which in the example would result in a left-right score of the cabinet of 4.57 ($4 * \frac{60}{140} + 5 * \frac{80}{140}$).

After calculating this score for all cabinets of the 12 destination countries from 2002 to 2014, we split the distribution of all cabinets in the sample at the median and code the cabinets below the median left and the cabinets above the median right. With the median cabinet left-right score in our baseline sample being 5.86, the cabinet in the above example would, thus, be coded as left.

When coding the quarters before and after the election, we account for the fact that in most cases of early elections, the date of the election is only announced shortly before the election. If an early election, for example, takes place in April and is announced in February, we would only code the quarter right before the election quarter as before the election.

In order to account for the political and economic situation in the source countries, we follow Hatton (2016) and add data on the Political Terror Scale (Gibney et al., 2016), the Freedom House Index of Civic Liberties and Political Rights (Freedom House, 2017),

⁷Note that only cabinet changes where the cabinet group switches between left and right are marked in the graph. There are many more small cabinet changes which do not cause a change in the cabinet position group.

the number of civil war battle death as measured by the Uppsala Conflict Data Program (Uppsala University Department of Peace and Conflict Research, 2017) and real GDP per capita from the Penn World Tables (Feenstra et al., 2015). With the exception of the number of civil war battle deaths, these variables only vary at the yearly level. To approximate the situation in the origin country at the time leading up to the decision to leave the country and to apply for asylum in another country, we use the average values of origin specific variables from the current and the 3 previous quarters. The political terror scale in quarter 1 in 2005, for example, is consequently calculated as the mean of the political terror scale in 2005 and 3 times the political terror scale in 2004.

To further capture the economic attractiveness of the destination countries we use data on the quarterly unemployment rates and the quarterly real GDP per capita in the destination countries from Eurostat. Finally, in some specifications we also include time-invariant bilateral information on the distance between origin and destination countries (Gleditsch, 2017) and the number of adult immigrants from the origin country living in the destination country in 2000 (Artuç et al., 2015). Table 3 provides some descriptive statistics for the first-time asylum applications, the cabinet left-right score, as well as for the origin, destination and bilateral control variables.⁸

[Table 3 about here.]

3 Econometric specification

In our main specification we estimate the following fixed-effect regression:

$$Y_{ijt} = \alpha_1 \mathbf{O}_{it} + \alpha_2 \mathbf{D}_{jt} + \alpha_3 [\mathbf{Q}_j \cdot \mathbf{C}_{jt}] + \tau_t + \sigma_{ij} + \varepsilon_{ijt}, \quad (1)$$

where the dependent variable (Y_{ijt}) represents the log of the number of first-time asylum applications per capita from citizens of origin country i in destination country j at time t . To derive this equation, we closely follow the specification of Hatton (2016), and add to the origin and destination specific explanatory variables \mathbf{O}_{it} and \mathbf{D}_{jt} , a set of interaction terms of the ruling cabinet's position indicator \mathbf{C}_{jt} and a set of dummies for before and after the election ($\mathbf{Q}_j := Q_{j,bef}, Q_{j,aft}$) or for different quarters before and after an election in a quarter $q = 0$, in destination country j ($\mathbf{Q}_j = \{Q_{jq}, q = -6, \dots, +6\}$).

Similar to Hatton (2016), we include Political Terror Scale, Freedom House Index of Civic Liberties, Freedom House Index of Political Rights, number of battle deaths and log real GDP per capita in the vector of time variant origin specific variables \mathbf{O}_{it} . The vector of time variant destination variables \mathbf{D}_{jt} comprises of the quarterly log real GDP per capita and the quarterly unemployment rate at destination. In order to control for the

⁸More detailed information on the data sources, as well as on the definition and the calculation of the individual variables can be found in the data section in the online appendix.

possible impact of the inflow of asylum seekers on the outcomes of asylum decisions, we additionally include the log of the yearly total asylum decisions in the destination country as well as the log of the yearly dyadic asylum decisions in the destination country in our regression equation (1). Moreover, our main specification as described in equation (1) includes quarter-year dummies τ_t and destination-origin fixed effects σ_{ij} . In all regressions standard errors are clustered by origin country.

4 Results

Table 4 presents the results of the estimation of equation (1) for the outcome “log first-time asylum applications per capita” for the model with only two periods: before the election comprising of the 5 quarters before the election and the election quarter, and after the election comprising of the 6 quarters after the election. As evident from Table 4 our results are robust to different specifications of the fixed effects: all coefficients are very similar to our main specification with destination-origin fixed effects and time dummies (column 2); both the version with origin fixed effects and destination and year dummies (column 1) and the version with origin-time fixed effects and destination dummies (column 3).

[Table 4 about here.]

Most importantly, our results establish empirically that asylum policies are affected by the electoral cycle and the identity of incumbent parties. As illustrated in Figure 2 we find that in the time before an election the inflow of refugees is very similar across all types of cabinets, whereas in the quarters just after an election, the inflow of refugees diverges substantially, with significantly less asylum applicants under right-wing cabinets and significantly more asylum applicants under left-wing cabinets.

[Figure 2 about here.]

The results of the estimation with individual dummies for different quarters before and after an election, which are illustrated in Figure 3, confirm that the turning point is really the quarter following the election. As highlighted in the next section this interesting pattern is robust to several different specifications and suggest that both left and right-wing cabinets choose moderate policies before the election and less moderate policies after the election.

[Figure 3 about here.]

In line with the literature and the results of Hatton (2016), we find that measures of political oppression and violence in the host country are positively correlated with

the number of asylum applications. The negative significant coefficient of the log origin country real GDP per capita suggests that also adverse economic conditions in the host countries drive asylum applications. As bad economic conditions are, however, often a byproduct of wars and political instability, this does not mean that asylum seekers leave their home country primarily for economic reasons. Moreover, we find a small negative effect of the unemployment rate in the destination country, which could indicate both, that a higher unemployment rate reduces the attractiveness of a destination country or that in times of higher unemployment more restrictive asylum policies are implemented. Interestingly, in contrast to Hatton (2016), our results (which are based on a different sample of countries) indicate that a higher GDP per capita in the destination country is associated with fewer asylum applications.

An important question is whether the uncovered patterns can be interpreted as causal effects. On this matter, a few comments are in line. One concern is that refugee inflows might have influenced the date of the elections, which took place outside the regular electoral cycle. Between 2002 and 2014 in most countries, however, migration was not one of the top issues on the political agenda. And indeed when researching the causes for individual incidences of early elections we did not find a single case, where migration in general or the refugee inflow in particular played a role for the decision to call early elections. In this sense, the estimates are likely to identify the causal effect of the electoral period on the admission of refugees. Second, there is a concern that previous refugee inflows would have affected the outcome of the election, biasing the results through an omitted variable problem. Reassuringly, controlling for previous levels of refugee inflow at the country level does not substantially change our results. Third, there is a question of whether unobserved short term shocks that affect the inflow of asylum seekers might influence the outcome of the election.⁹ In this case, the cabinet position is endogenous and the estimates can be understood as upper bounds to the true effect of a given party on the refugee influx.

Finally, a deeper question is whether the effects of elections and political parties on refugee inflows are driven by the demand side (the refugees) or the supply side (the incumbent party). As emphasized by the economic literature, asylum seeking behaviors are mainly driven by exogenous factors in the home country. It is difficult to imagine that conditional on the identity of the incumbent party, refugees wait for the outcome of the election to file or not their asylum application. However, asylum seeker might select into different countries. To believe this, one needs to assume that asylum seekers have relatively good knowledge of the current political situation in the receiving country. Since we consider short-term quarterly variation of asylum flows, we rather believe that the effects captured are mainly driven by the supply side.

⁹A recent example is the Green Party in the Netherlands, that refused to take part in a majority coalition because of the stance of other parties on refugee policies (*The Economist*, 2017).

5 Robustness Checks

Our robustness checks (available upon request) show that our main results are very stable across a large number of different specifications of our regression equation (1). Among others, we control for the average asylum applications per capita in the previous 5 years, which indicates that election outcomes are not influenced by previous refugee inflows. A further concern is that some of our effects might be driven by the change in the data collection method by Eurostat in 2008. In order to account for that, we therefore also conduct a robustness check including a dummy which is equal to 1 if the year is 2008 or later and equal to zero if the year is 2002 to 2007. As the coefficient of this dummy is small and insignificant and the results do not change, we are confident that change in the collection method does not influence our results. Several other robustness checks, moreover, show that our results are not sensible to

- using the log of the number of first-time asylum applications per capita in the in the **origin** country as our dependent variable
- using the current values of the origin country control variables instead of averages of the current and the past three quarters
- clustering the standard errors on the destination-origin level
- normalizing the cabinet position on the country level before computing the cabinet position dummies (left and right)
- leaving missing data on quarterly first-time asylum applications for 2008 and 2009 missing instead of imputing it from the total applications
- choosing different cutoffs for dropping countries pairs with few applications (one and three average applications per quarter)
- looking only at five or four quarters around the election

Finally, we test whether our results depend on the sample of countries used. The results only change marginally when we include Cyprus in the sample of destination countries. Moreover, when adding very small countries in terms of first-time applications (Estonia, Latvia, Lithuania, Portugal, Malta and Slovenia), the main pattern is still visible in the data.¹⁰

6 Conclusion

We examine the impact of elections and parties on first-time asylum applications based on a large bilateral panel data set comprising 12 European destination countries and

¹⁰Note that for each sample of destination countries we adjust the sample of origin countries such that in total the origin countries account for more than 90 percent of all first-time applications in the destination countries between 2002 and 2014.

their 51 most relevant origin countries during the time period 2002 to 2014. Our findings suggest that the number of asylum applicants under left- and right-wing parties converges before elections and differs thereafter. In conclusion, our results clearly show a strong impact of elections and parties on first-time asylum applications. This highlights the need to better model the influence of political economy factors such as elections or interactions among receiving countries when analyzing the determinants of refugees inflows (Görlach and Motz, 2017).

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Figure 1: Elections and cabinet changes

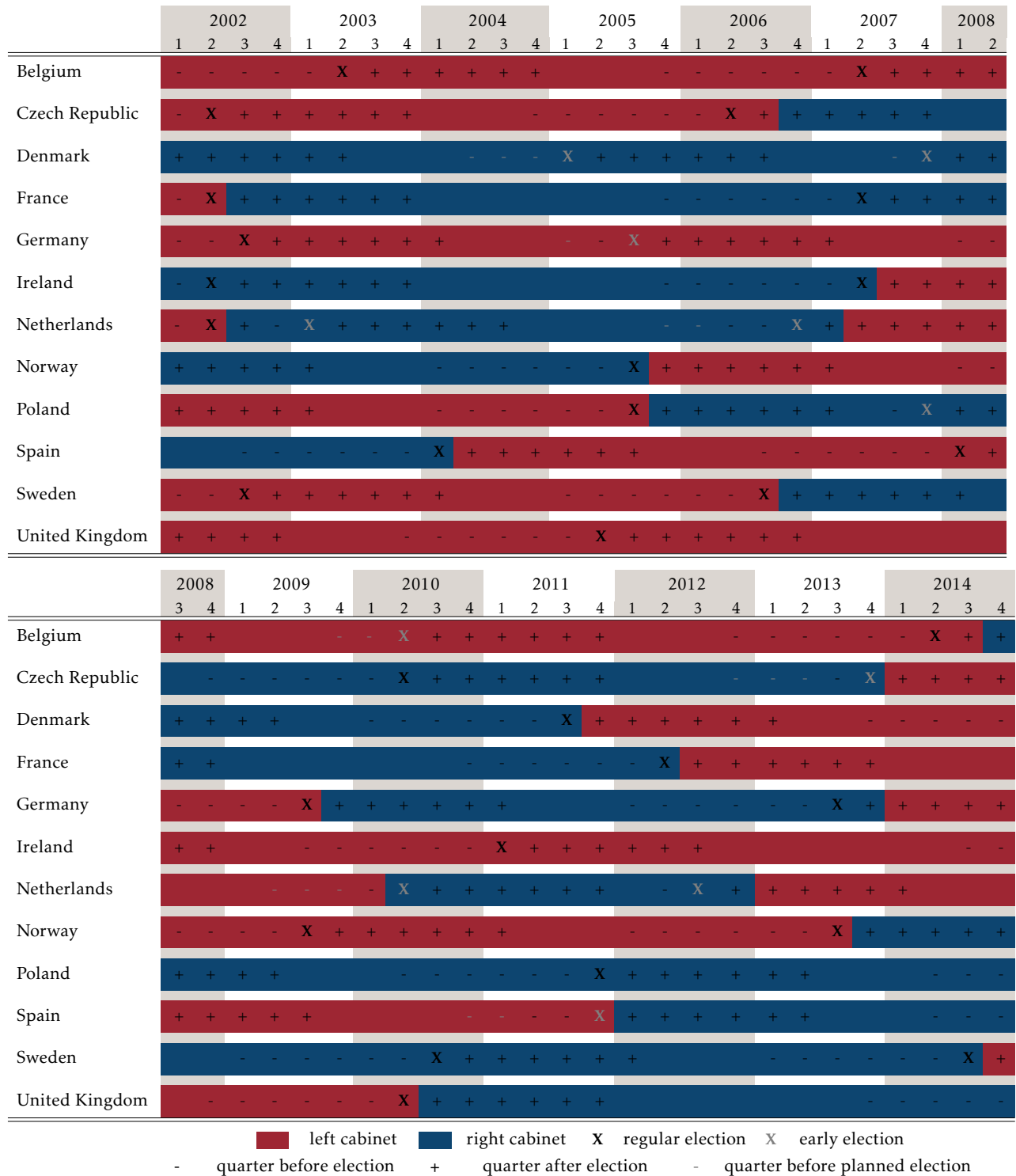
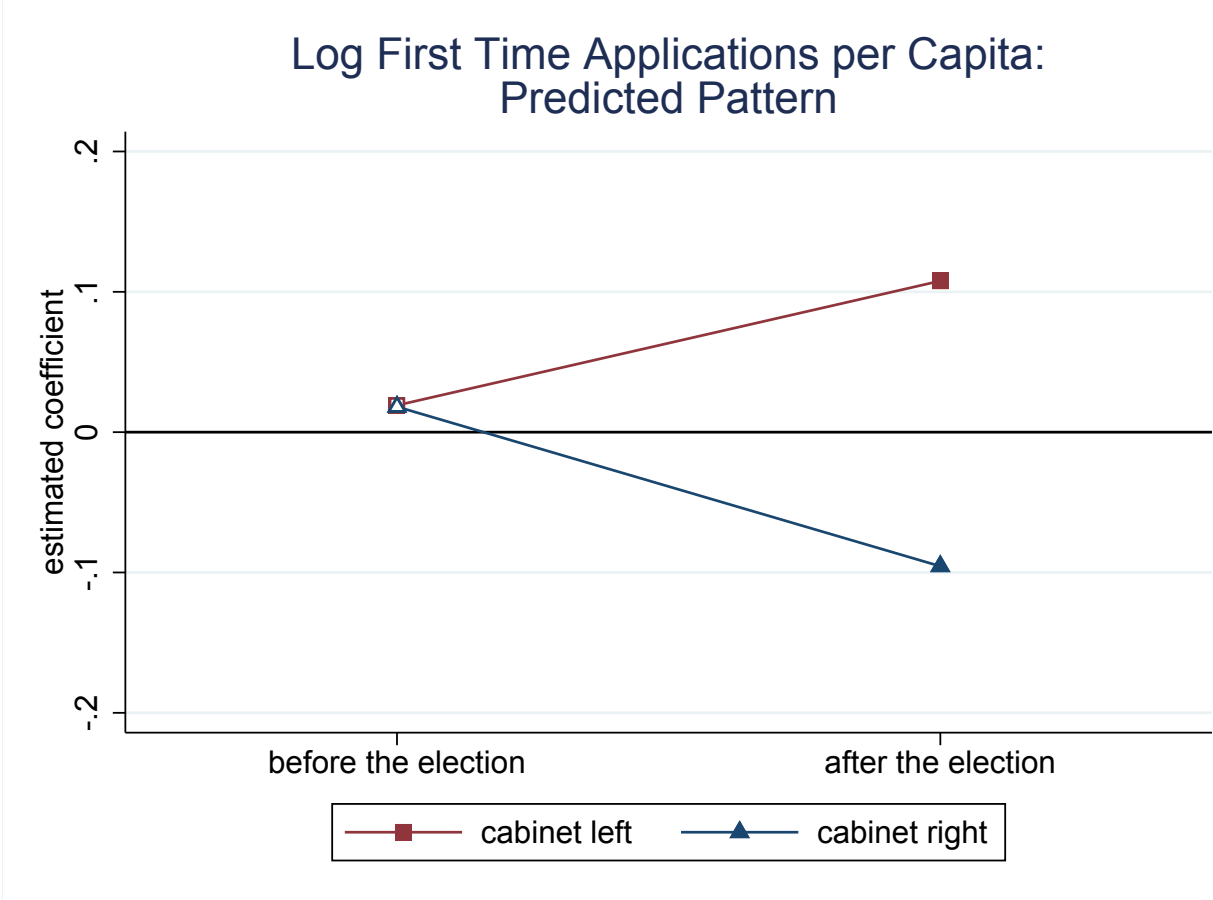
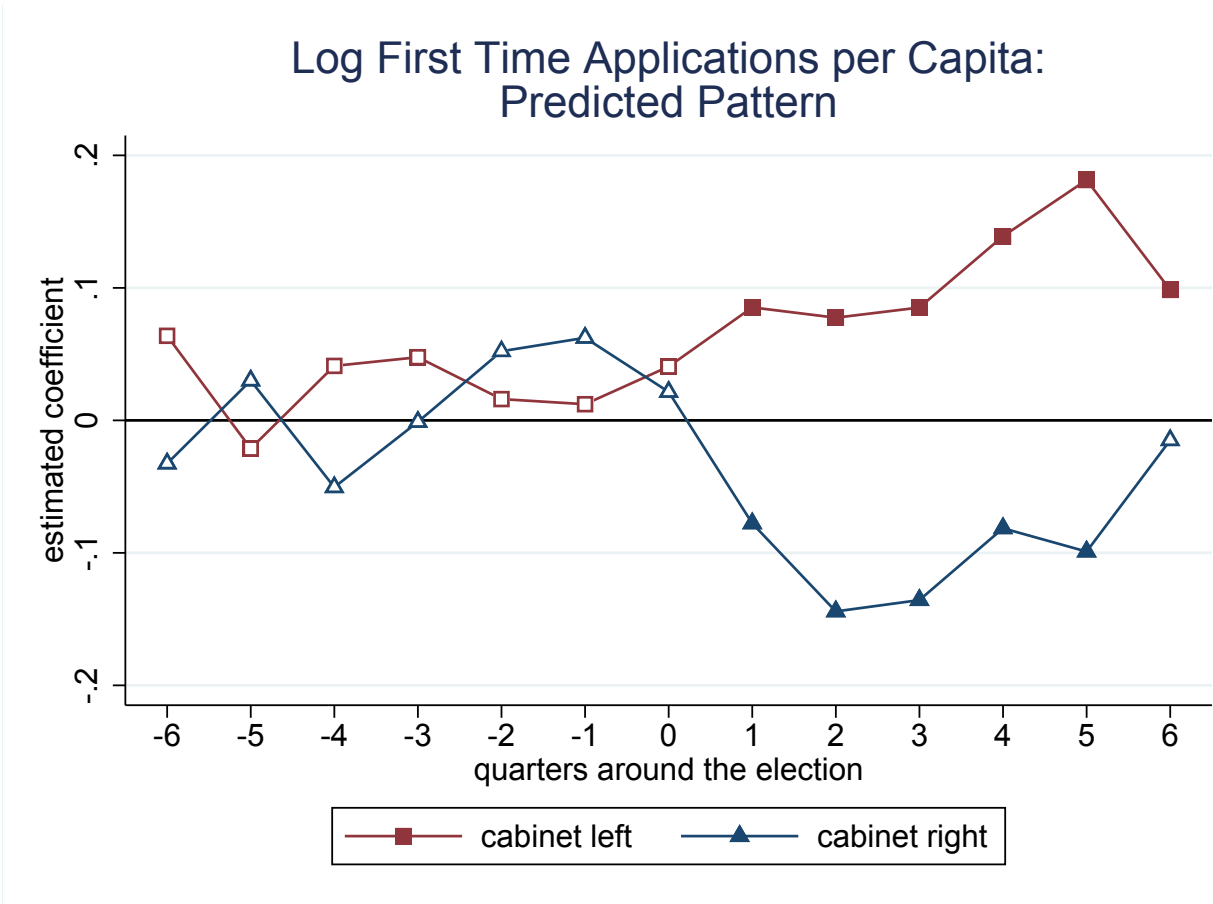


Figure 2: Log First-Time Asylum Applications per Capita: Predicted pattern before and after an election



This figure shows the time evolution of refugee inflows as estimated in the fixed effects regression (1) with a set of dummies for before and after the election. Significant coefficients are indicated by filled plot markers.

Figure 3: Log First-Time Asylum Applications per Capita: Predicted pattern 6 quarters before and after an election



This figure shows the time evolution of refugee inflows as estimated in the fixed effects regression (1) with a set of dummies for different quarters before and after an election in a quarter $t = 0$. Significant coefficients are indicated by filled plot markers.

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Table 1: Total number of first-time asylum applications between 2002 and 2014 in 12 European destination countries

Destination country	# of first-time applications
Germany	704450
France	629288
United Kingdom	470960
Sweden	445525
Belgium	184200
Netherlands	167055
Norway	113545
Poland	89680
Denmark	59440
Spain	56227
Ireland	47070
Czech Republic	35370

Notes: The number of first-time applications represent the sum of first-time applications in all available quarters from Quarter 1 2002 to Quarter 4 2014. For France the number of first-time applications in 2008 and for Spain the number of first-time applications in 2008 and 2009 are imputed from the number of origin-specific applications in these years. For Belgium no data is available in 2004 and for Norway no data is available in 2002 and in Quarter 2, 3 and 4 of 2007.

Table 2: Top 10 source countries and their share in the total first-time applications in the 12 destination countries from 2002 to 2014

Source country	Share of first-time applications	Cumulative share
Russia	7.0%	7.0%
Iraq	6.9%	13.9%
Syria	6.1%	20.0 %
Afghanistan	5.2%	25.2%
Somalia	4.6%	29.8%
Iran	3.4%	33.1%
Turkey	3.4%	36.5%
Eritrea	3.3%	39.8%
Serbia	3.0%	42.9%
Democratic Republic of Congo	2.8%	45.6%

Table 3: Summary statistics

Variable	Obs	Mean	Std. Dev.	Min	Max
Quarterly fist-time asylum applications	23705	114.18	338.11	0	15330
Quarterly first-time asylum applications per 100,000 inhabitants	23705	.57	2.18	0	112.55
Number of elections per destination country	23705	3.45	.81	2	5
Number of cabinet changes per destination country	23705	1.83	.87	1	4
Left-right position of the cabinet	23705	5.57	1.52	2.77	8.22
Political Terror Scale	23705	3.34	.93	1	5
Civic Liberty (FHI)	23705	4.58	1.45	2	7
Political Rights (FHI)	23705	4.87	1.7	1	7
Quarterly civil war battle death (000s)	23705	.22	.87	0	15.09
Yearly real GDP per capita at origin	23705	6440.79	5270.3	336.8	24039.13
Distance from origin to destination	23705	4395.18	2167.79	454	9680
Migrant stock in 2000/1	23705	16452.49	74737.36	0	1272000
Quarterly real GDP per capita at destination	23705	8718.8	3206.1	1557.45	18047.84
Quarterly unemployment rate at destination	23705	7.76	3.93	2.4	26.9

Table 4: Determinants of log(First time asylum applications per capita)

	(1)	(2)	(3)
Political Terror Scale	0.399*** (0.0703)	0.400*** (0.0697)	
Civic Liberty (FHI)	0.174 (0.132)	0.175 (0.131)	
Political Rights (FHI)	0.0472 (0.0753)	0.0486 (0.0752)	
Quarterly civil war battle death (000s)	0.190*** (0.0236)	0.188*** (0.0234)	
Log origin country real GDP per capita	-0.659*** (0.165)	-0.661*** (0.163)	
Log migrant stock in 2000/1	0.263*** (0.0210)		0.263*** (0.0210)
Log distance from origin to destination	-0.608* (0.298)		-0.613* (0.296)
Log destination country real GDP per capita	-1.404** (0.490)	-1.479** (0.441)	-1.146* (0.465)
Quarterly unemployment rate at destination	-0.0731*** (0.0114)	-0.0743*** (0.0108)	-0.0715*** (0.0116)
Cabinet position left * Before the election	0.0207 (0.0262)	0.0191 (0.0261)	0.00933 (0.0269)
Cabinet position left * After the election	0.116*** (0.0228)	0.108*** (0.0212)	0.111*** (0.0229)
Cabinet position right * Before the election	0.0159 (0.0249)	0.0181 (0.0234)	0.0191 (0.0247)
Cabinet position right * After the election	-0.101*** (0.0242)	-0.0955*** (0.0230)	-0.101*** (0.0240)
Observations	23705	23705	23705
Adjusted R^2	0.444	0.176	0.447
Fixed Effects	O	D x O	O x T
Destination dummies	Yes	No	Yes
Quarter-Year dummies	Yes	Yes	No

Standard errors in parentheses

* $p < 0.05$, ** $p < 0.01$, *** $p < 0.001$

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