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# Nocturnal satellite imagery and the analysis of spatial change

Rolf Bergs<sup>1</sup>

## 1. Introduction

Night satellite images may offer an interesting tool to generate socio-economically relevant data and to analyse the evolution of space, e.g. cities or rural areas, and how spatial units interact over time. This paper is just an essay with preliminary ideas for discussion; the approach is explorative-methodological rather than one putting an empirical focus on a defined research item. Empirics discussed in this paper are just various examples collected from tinkering with image analysis and statistics software.

Henderson et al. (2011) found that there is a significant association between economic growth and the changing emission of visible light. Mean correlation coefficients between luminosity and GDP (PPP) found by Chen and Nordhaus (2011) are stable at slightly more than 0.8 with a minor standard deviation of just 0.009 (estimated from data in: Chen and Nordhaus 2011, annex, p.11). The connection between both variables is energy consumption for public utilities, communication, production and settlements that all needs light. The advantage of visible light emission (in the wavelength range between 380 - 700 nm) is its one-dimensional scope of data generated: There is only a range between black, the different grey shades and white, while daytime images give one some idea of settlement, commercial, green and traffic zones with a high variation of colours of different brightness that are hard to translate into meaningful raw data.

A major purpose of using night satellite images for economic analysis has been the search for proxies for production and population density in countries with insufficient and unreliable data infrastructure. This essentially applies to less developed countries where weak data infrastructure is often part of overall underdeveloped administrative capacities. Hence, there was the idea to estimate important social and economic profiles of such countries and regions by the luminosity proxy and to improve the decision base for policies and aid. <sup>2</sup>

The fundamental assumption is that there is not only a significant empirical relationship between luminosity and income (or labour productivity respectively) and luminosity and population, but that error in both, luminosity and socio-economic data, is independent. The latter assumption is even more important than the former:

“...In general, a positive correlation between measured income (national accounts or survey means) and nighttime lights could be due to two factors: that they are both correlated with true income, or that their measurement errors are strongly correlated with each other. However, the latter possibility is implausible because the generating process of nighttime lights data is to a very large degree independent of the

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<sup>1</sup> PRAC – Bergs & Issa Partnership Co., Bad Soden, Germany: The draft paper was presented at the 16<sup>th</sup> ERSA Congress, Vienna, August 23 - 26, 2016 (earlier draft versions were titled: “Exploring the spatial economy by night”)

<sup>2</sup> The ultimate motivation for this paper stemmed from a brief introductory paper of the National Optical Astronomy Observatory (NOAO) GSMT programme. One component of this programme deals with measuring light pollution. The advantage of ImageJ as a powerful image analysis software is described (<http://www.noao.edu/education/gsm/Ipmeasure>).

generating process either of national accounts or of survey means. For example, measured income is collected by statisticians interacting with survey respondents, while nighttime lights are recorded impersonally by satellites..." (Pinkovskiy and Sala-i-Martin 2014, p. 2).

Light emission is essentially a by-product of production and consumption (cf. Henderson et al. 2011). The brightness of nocturnal light reveals the density of human activities. Error variance of luminosity (spread of the difference between measured and real light emission) is rather constant around the globe while there is a large variation of error in datasets of official social and economic statistics. In many developing countries error in official statistics is larger than that of light emission, while in industrialised and newly industrialised countries official data are more reliable so that further satellite data on luminosity would not add value to statistical information in those countries. That is the reason why night satellite analysis is of relevance for developing countries rather than for industrialised ones (Henderson et al. 2011, Chen and Nordhaus 2011). However, this differentiation only holds for purposes to derive proxies for production or population data.<sup>3</sup> In fact, there is also reason to use this tool in the observation of (spatial) economic patterns and trends in the more industrialised countries. A major advantage is the observation of change in a comparison of images. In this case, measurement error (see below) is of less relevance since it is cancelled out to a large extent. Evolution of spatial distribution of rural areas, urban agglomerations, border areas or other spatial categories are to be mentioned. My hypothesis is that several indicators, such as changing border effects, growing peri-urban zones, trends of polycentric or monocentric development or the rank size distribution of settlement clusters can be made visible and further processed by using night satellite images. Patterns of spatial dependence and spatial heterogeneity are thus at the centre of this study.

Just comparing normal geographical maps is insufficient to shed light on urban and rural change. Enriching topographic or political maps with external data (GIS) helps to monitor spatial change by variables for which data are available. This is, however, only possible at a higher administrative level, such as countries, regions and - to a limited extent - districts. Thus, by using official statistics the view on space is truncated. Micro-spatial patterns and processes cannot be observed that easily, even though there is a growing literature on using micro-spatial raster data for the research on neighbourhood effects viewed from different perspectives, such as labour market, education, migration, investment choice and others (e.g. Halleck Vega and Elhorst 2014, Zwiers et al. 2014, Coulter et al. 2015, Duranton and Overman 2005). The advantage of nocturnal satellite images is the global access at any spatial scale without the need to refer to administrative entities and boundaries. The purpose of this paper is to explore the information content of the nocturnal satellite images for the analysis of patterns and change of the economy at any spatial scale. The tool examined for this exploration is ImageJ, an open but powerful image analysis package.

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<sup>3</sup> Even though there is less association between economic output and luminosity in areas of lower light emission density, a specific property of large backward areas in the developing world (Nordhaus and Chen 2015, p. 218 f.).

## 2. *Spatial heterogeneity, spatial dependence*

The night satellite images may provide an opportunity to explore patterns of spatial dependence (e.g. how space is influenced by further contiguous or distant space) and spatial heterogeneity (e.g. how attributes of space are distributed, whether there are regular or irregular patterns). In econometric modelling with a spatial relation, both features are of major relevance to consider, when it comes to efficient estimates. The violation of two *Gauss-Markov* assumptions is to be acknowledged, namely the constant variance of variables over space and the absence of covariance between observations at different places. Consequently, heterogeneity and dependence can be conceived also as stand-alone categories to describe and analyse space. Spatial dependence and spatial heterogeneity are both results of processes; while heterogeneity is the result of a deterministic process and cause of a further continuing process, spatial dependence is rather a stochastic process in which spatial variables in one territorial area are influenced from other contiguous or distant areas. In practice, spatial dependence and heterogeneity can be conceived as follows:

Spatial heterogeneity is related to the variation of relationships over space. It is thus a property describing the uneven distribution of variables describing space, such as population, topography, economic assets etc. If the world were a fully homogenous space, there would be no need for transport because every point of space would be identical such as in a closed vacuum. Cities side-by-side with rural areas, different levels of population density, converging and diverging regions or simply maritime versus continental climate areas are indications of spatial heterogeneity. But the concept of spatial heterogeneity also comprises a more holistic view on spatial structures, often based on fractal geometry such as multi-scalar global patterns that come close to natural laws, like the *Zipf* law of rank-size distribution of cities or the forms of coast lines or topographic reliefs. The concept of “natural cities”, their detection by nighttime imagery and the subsequent estimation of their characteristic distributions is an example for that (Jiang et al. 2015).

Spatial dependence in a sample dataset means that observations on a variable  $k$  at a location  $i$  are not independent from observations  $k$  at other contiguous or more distant locations  $j \neq i$ . A central attribute of spatial dependence is the covariance of selected random variables over space: While spatial heterogeneity represents the first-order moments of space, spatial dependence is traditionally conceived as a second-order property.<sup>4</sup> Spatial units influence each other to a varying extent – primarily depending on distance, markets, politics and social fabric. An indicator of spatial dependence is spatial autocorrelation. In spatial econometric modelling it is considered either by spatial autoregressive or spatial error models (see footnote 6 below). But spatial autocorrelation can be also used as a stand-alone indicator for descriptive statistical analysis. This way, the *Moran-I* coefficient (in analogy to the *Pearson-r*) or the *Geary-c* coefficient are measures of estimating spatial dependence. (Anselin 2001, Le Sage 1999)

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<sup>4</sup> Even though, there are arguments that the relationship is contrary, namely that spatial heterogeneity is rather a second-order effect, while dependence is the first-order one. Jiang (2015) argues that “...Spatial heterogeneity is a kind of hidden order, which appears disordered on the surface, but possesses a deep order beneath. This kind of hidden order can be characterized by a power law or a heavy-tailed distribution in general. ...”. So if just viewing the visible heterogeneity at the surface, then the dependence between spatial items is a higher order effect, but if looking at underlying regular patterns of heterogeneity, spatial dependence constitutes a lower (first) order property.

### 3. *Further literature*

Meanwhile there is growing literature dealing with economic research based on night imagery. Henderson et al. (2011), Chen and Nordhaus (2011) and Sutton et al. (2007) primarily deal with the examination of light emission as a proxy for missing economic and social data (most importantly production output and population). Mellander et al. (2011) examine the explanatory power of night imagery specifically for economic activity rather than economic growth, population density or urbanization for Sweden. Ghosh et al (2010) used the night imagery to develop a method to identify informal economic activity that is not recorded in official statistics. Pinkovskiy and Sala-i-Martin (2015) examine whether GDP per capita recorded in national accounts or mean income and consumption data from household surveys are better proxies for income in linking the evolution of both series to the change of night image luminosity. Pinkovskiy (2013) looks at discontinuities of the political economy along borders using a regression discontinuity design. Border discontinuities are supposed to shed light on the impact of the political economy on economic activity because these determinants are produced by government activity. Luminosity per capita and its change over time behaves discontinuously upon crossing a border from a poorer (or lower-growing) into a richer (or higher-growing) country. According to Pinkovskiy these discontinuities seem to form lower bounds for discontinuities in economic activity across borders, which suggest a major importance of national-level variables such as institutions and culture relative to local-level variables. In another paper, Chen (2015) used DMSP-OLS radiance-calibrated images from 2006 to examine infant mortality and poverty rates.

There are several further studies linking night imagery and socio-economic variables – the majority on developing countries - but there are only rather few (e.g. Jiang 2015, Jiang et al 2015; Liu 2014; Zhang et al. 2016) using the night imagery to shed light on spatial dependence and spatial heterogeneity, both of major relevance in spatial economics.

### 4. *Images*

The images analysed in this paper are satellite images of the National Oceanic and Atmospheric Administration (NOAA).

Global nighttime lights imagery data are collected by the Defense Meteorological Satellite Program Operational Line Scanner (the DMSP-OLS Nighttime Lights Global Composites - Version 4 1992-2010). Their original purpose has been the detection of moonlit cloud coverage. Several DMSP-OLS image categories that are useful for spatial analysis are freely accessible on the internet. The night images recorded since 1992 have a spatial resolution of 30 arc-seconds and are recorded between 75°N and 65°S with non-calibrated light intensity in a Digital Number (DN) range between 0 and 63 (6-bit). Light intensity across pixels is *Pareto*-distributed.

The DMSP-OLS is capable to record and observe light sources in a range between visible-near infrared emissions and strong luminosity, present at night on the Earth's surface, including settlements, roads, gas flares, ships and fires.

The available so-called “products” of imagery since 1992 are coverage, average visible band values, stable light and average light normalized.

In our examples below I mostly used the average digital number of lights multiplied by the percent frequency of light detection (*avg\_lights\_x\_pct*). These images are normalised for variations of cloud cover and thus account for the persistence of lights, but one can also test images with stable lights, average visible light or the additional calibrated images, that are available for selected years only. The 2011 radiance-calibrated image is included for comparison purposes.

### 5. Measurement error

In a static comparison of images, a large portion of error is cancelled out, since the factor of error is relatively stable. However measurement error increases with declining spatial scale and the level of measured light emission. Zero luminosity is only affected by error due to light below visibility.

The ordinary DMSP-OLS images do not indicate radiance but only a light intensity within a defined DN range between 0 and 63. The observations are thus right-censored. For heavily lit areas, such as city centers, the problem of saturation hampers precise measurement.<sup>5</sup> However, the distribution of real luminosity is highly skewed and only around 0.1 percent of the observations is topcoded (Henderson et al. 2011, p. 195).

For the product “stable lights” background noise (moonlight reflection etc.) is removed. Background noise of night imagery showing average luminosity can be filtered out by setting empirical thresholds based on visible light identified in areas with no source of light emission. Due to variation of light emission it is difficult to precisely filter out over-glow. If one looks at the distribution of light emission on the satellite images it is intuitively clear, that a bright area is not illuminated by light from this area alone but also beamed at from neighbouring space and *vice versa*. As Doll (2008) shows, “...a feature of the data acquisition process is that there is a large overlap (some 60%) between pixels. This means that light observed in one location has the chance to be recorded in more than one pixel. This can contribute to a larger lit-area being detected than is actually the case...” (p.12). In spatial regression analyses, average luminosity of a small area depends on light emitted in that area and – to some extent - from contiguous areas<sup>6</sup>. The larger the areas regarded, the less will be the spillover impact from over-glow.

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<sup>5</sup> Only recently there have been efforts to solve the saturation problem. A set of radiance-calibrated images is available for selected years only (cf.: [http://www.ngdc.noaa.gov/eog/dmsp/download\\_radcal.html](http://www.ngdc.noaa.gov/eog/dmsp/download_radcal.html))

<sup>6</sup> However, spill-overs are also to be considered for larger areas if visible luminosity  $m$  is further related to other variables, such as population, labour productivity or other predictors  $X$ , more formally:  $m = \rho Wm + \beta X + \varepsilon$ , where  $W$  is a weighting matrix defining contiguity or distance respectively and  $\varepsilon$  the error term in a spatial autoregressive model. In an alternative spatial error model the specification would be:  $m = \beta X + \varepsilon$ , where  $\varepsilon = \lambda W\varepsilon + v$

Apart from that it is important to stress the fact that light emission in its association with economic variables is also affected by further sources of error. Comparing population density across countries by using the luminosity proxy will be misleading in many cases, simply because of lower levels in light emission in developing or emerging economies. Ghosh et al. 2010 point out that e.g.

“ ... countries such as Japan and Germany are much wealthier relative to their sum of lights value; on the other hand, Russia is much poorer relative to its sum of lights value. ...” (p. 153)

At micro- or meso-scale specific land use of economic activity may have an influence on light emission after dark. Henderson et al. 2011 mention the example of the cities Las Vegas vs. Salt Lake City, or the fact that high density multi-storey buildings emit relatively less light compared to flat buildings. Hence, even though there is a significant correlation between luminosity and production over space, it would be rather misleading to use luminosity as a general proxy for missing GDP data, especially for partial analyses with lower numbers of observation.

In addition, different satellites with different technology levels may generate images that are not completely comparable over time, just as a sampling variation due to physical factors. The available OLS images are not inter-calibrated. Inter-calibration of images is important for time series analyses, where systematic breaks through changes of satellites and technology distort the data. In this paper I just compare images at two distinct points of time, hence for the purpose of simplicity, in this paper the simple use of the raw OLS images is deemed justified<sup>7</sup>.

## 6. *The Software ImageJ*

Standard of spatial analysis of maps and geographic images is ArcGIS, a powerful, costly but also rather complex software for geographical analysis, including image analysis. For the analysis of night imagery, pure image analysis software is in most cases sufficient. ImageJ is a widely used open source image analysis software that had been developed by Wayne Rasband at the National Institutes of Health since the end of the 1990s (Schneider et al. 2012). Its primary purpose has been to process images in medical and biological sciences where has obtained a *de facto* standard for image analysis. However, ImageJ as a processing and data analysis toolbox, written in Java, has a broad range of applications - beyond microscopy and medical imaging - such as astronomy<sup>8</sup>, environmental analyses<sup>9</sup>, earth sciences including remote sensing<sup>10</sup>, material sciences<sup>11</sup>, viticulture<sup>12</sup> archaeology<sup>13</sup> and others. The software has

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<sup>7</sup> Chen and Nordhaus (2011, p.3) have estimated the standard errors of visible light emission on the DMSP-OLS images by regressing luminosity on GDP. Mean standard errors of estimates for luminosity between the four different satellites, on the one hand, and for one satellite over years, on the other hand, are found in the range between 0.3 and 0.6 with a decreasing trend over time for 1°x1° grid cells (Chen and Nordhaus 2011, annex, p. 9). A large inter-calibrated national database can be found on [http://ngdc.noaa.gov/eog/dmsp/download\\_national\\_trend.html](http://ngdc.noaa.gov/eog/dmsp/download_national_trend.html). See also: Eldvidge et al. (2014) and Liu (2014) .

<sup>8</sup> West and Cameron 2006; Damas et al. 2014

<sup>9</sup> E.g. van Woesik et al.; Troscianko and Stevens 2015

<sup>10</sup> E.g. Berra et al., Buntlow 2013

<sup>11</sup> E.g. Ushizima et al

<sup>12</sup> Whalley and Shanmuganathan 2013

been permanently further developed. The most powerful feature of ImageJ is its international community with innumerable contributions consisting of plugins and macros. Several tools originally developed for medical imaging are also useful for remote sensing and the analysis of spatial change.

ImageJ is capable to display, print, edit, analyse, process and save images with 8-bit, 16-bit and 32-bit graphic formats. It can read various file formats, such as tiff, bmp, jpg, gif and others. The software allows calculating area sizes and performs a large range of statistical and mathematical procedures on pixel values and their distribution either globally or in defined selections of any shape. With numerous processing functions such as contrast manipulation, sharpening, smoothing or filters, it is possible to visualize relevant parts of the images. Density histograms and profile plots are examples of statistical diagrams generated. Basic parameters of statistical analysis are mean values and spread but it is also possible to run more sophisticated analyses starting from image correlation to detecting fractal structures. It is easy to import the numeric data in statistical packages like R or Stata in order to merge datasets and to run further statistical or econometric procedures. The package is therefore also widely used as a remote-sensing software. Distances and angles can be measured. Even though the software is not geo-sensitive itself it is possible to calibrate geographical images containing geo-reference by figure calibration.<sup>14</sup> In the numeric dataset of moments the geographic coordinates can be saved and further processed.

## 7. *Spatial analysis with ImageJ*

### 7.1 *Selection*

ImageJ comprises a variety of area selection tools. The simpler ones are point and multi-point, rectangular, elliptical, polygon, freehand and straight line selections. These are part of the core software with button commands or commands from the pulldown menu. However, there are several further more complex selection tools provided as plug-ins or macros. Powerful plugins are e.g. the application of concentric circles or the radial profile. It is also possible to write macros for different individual kinds of selection (see below).

### 7.2 *Partial analysis*

A major advantage of the night satellite images is independence from official data and administrative boundaries. It is possible to define the partitioning of a map at the discretion of the researcher. The easiest tool of partial analysis is simply to crop the map and to focus the analysis on a defined section. Within the defined section it is then possible to select further sub-parts. In using the command “restore selection” the selection of an image at  $t_0$  can also be copied to the same image at  $t_n$ . In comparing results of alternative partitioning it

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<sup>13</sup> Verhoeven et al. 2013

<sup>14</sup> Geographic images are usually geo-referenced in a geo-tiff format. Even though ImageJ itself does not process and save the geo-referencing information it is possible to calibrate the image with the frame coordinates by the plugin “Figure Calibration”. As long as the image is open, any analysis – also of parts cropped to a new image file – can be done. The subsequent numeric analysis can be collated to the frame coordinates.



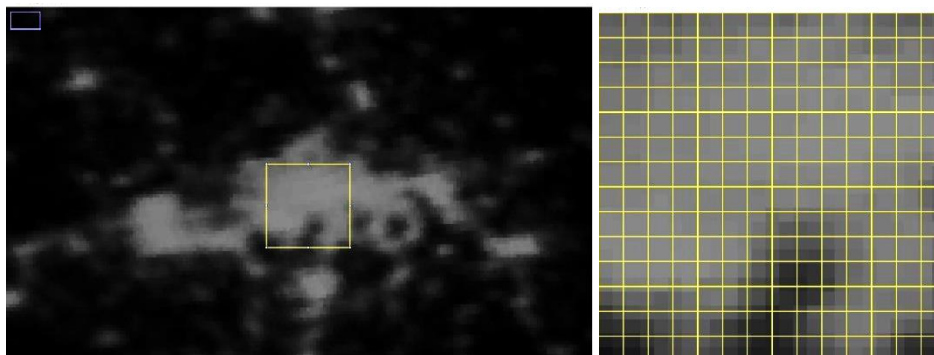
is possible to run various (infinite) sensitivity analyses. This can be very interesting for analyses of spatial change and spatial dependence such as spatial autocorrelation analyses. One possible approach could be to select sub-regions by free-hand (e.g. according to administrative boundaries); another approach can be to subdivide a map into raster cells, e.g. square kilometers<sup>15</sup>. This can be easily realized by writing simple ImageJ macros like the following one:

```
x = 0;
y = 0;
width = 1.785714286;
height = 1.785714286;
spacing = 0;
numRow = 100;
numCol = 100;

for(i = 0; i < numRows; i++)
{
    for(j = 0; j < numCol; j++)
    {
        xOffset = j * (width + spacing);
        print(xOffset);
        yOffset = i * (height + spacing);
        print(yOffset);
        makeRectangle(x + xOffset, y + yOffset, width, height);
        roiManager("Add");
        if (roiManager("count") > 10000)
        {
            print("Maximum reached: 10000 entries have been created.");
            exit;
        }
    }
}
```

Moments of such raster cells (either in total or selected cells) can be further processed and subsequently analysed by statistical procedures, either by ImageJ itself or statistics software. The smaller the cells the more it is possible to delineate administrative units by raster cells. It is then possible to add up different raster cells and compare the selected units over time.

The following example shows a selected raster (1x1 kilometer cells) of the Frankfurt region taken from the 1992 image:



Further below in this chapter examples of different selection procedures and statistical analysis are discussed.

<sup>15</sup> The maps of DMSP-OLS are scaled to one pixel = 0.56 kilometers

### 7.3 Visual inspection

One, perhaps even the most powerful tool of the maps is simply the image itself. The following night satellite image with a medium transparency shows the Korean peninsula. The northern and the southern part are immediately visible by a discontinuity of light emission exactly along the political border between both.



Satellite image from: DMSP-OLS Nighttime Lights Global Composites (Version 4): Snapshot of the Korean peninsula in 2010

In many cases, the information provided by the first glimpse at the image already contains the main message. Statistical analysis is then only needed to verify or falsify the estimate from visual inspection.

### 7.4 Statistical analysis of images

Comparative analysis of night images may shed light on spatial change, like urbanization or economic integration processes. The former is visible around cities, while the latter notably along administrative boundaries like national borders. With a simple image correlation analysis it is possible to view the variance of the pixel-by-pixel DN at two or more points of time. A high correlation coefficient (*Pearson's r*) will indicate either a little or a very balanced change, so that all pixels change their intensity with roughly the same rate. If there is a stronger urbanization process over time, image correlation will be lower, because there will be a certain portion of black pixels becoming grey or white while also many formerly black pixels remain black. Then the more the regarded city is zoomed in, the lower the image correlation coefficient will be.

More insight into patterns of spatial dependence will be provided by the analysis of spatial autocorrelation. Visual inspection of the night images reveals that both, areas with high and higher luminosity as well as those with zero and low luminosity are clustered and that between both types of areas there is smooth transition. It is unusual that pixels with DN=63 are side-by-side with zero pixels<sup>16</sup>. This shows that spatial units are not independent from neighbour units. However, the level of global spatial dependence of an area or local spatial dependence between neighbouring units within an area may change over time. Areas close to a city that were formerly rural may have attracted investment into production or housing showing a peri-urbanisation process that is indicated by higher luminosity, similar to that of the city itself. Rural areas become functional for the metropolitan area of a city, and spatial dependence will thus increase. Such a process can be indicated by a comparative analysis of spatial autocorrelation using the *Moran's I* coefficient.

$$I = \frac{N}{\sum_i \sum_j w_{ij}} \cdot \frac{\sum_i \sum_j w_{ij} (X_i - \bar{X})(X_j - \bar{X})}{\sum_i (X_i - \bar{X})^2}$$

where  $N$  is the number of observations and  $w_{ij}$  are the elements of the weight matrix.

In an evolution of functional areas increasing spatial autocorrelation would indicate higher integration of urban and former rural areas, especially if the number significant local coefficients (*Moran's local I*) increase. Declining spatial autocorrelation would indicate the opposite. Further below some examples of such analyses are illustrated.

In the following chapters there is a discussion of some examples ranging from visual inspection to further statistical analysis of images, such as spatial autocorrelation.

## 8. *Urban evolution from the macro perspective: Urbanisation and polycentricity vs. monocentricity*

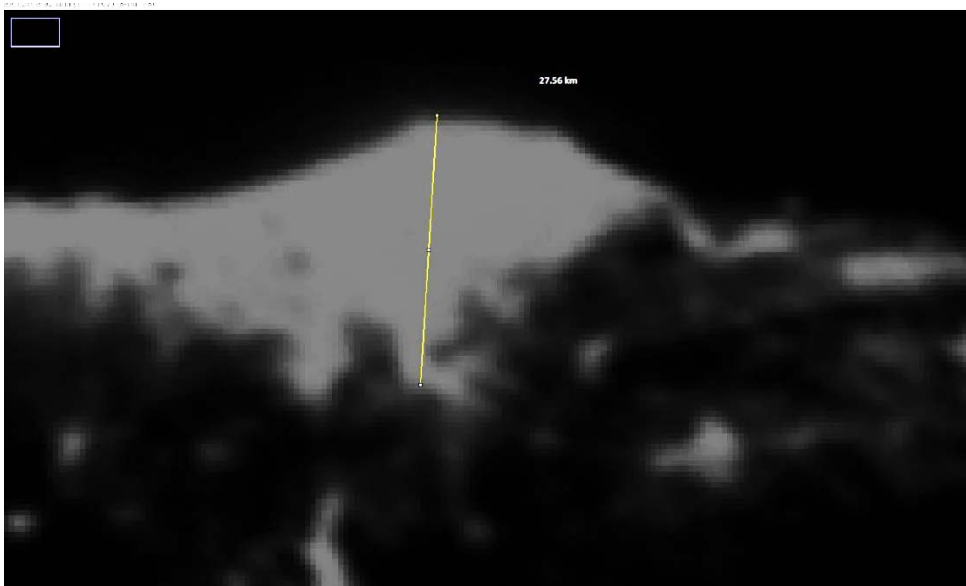
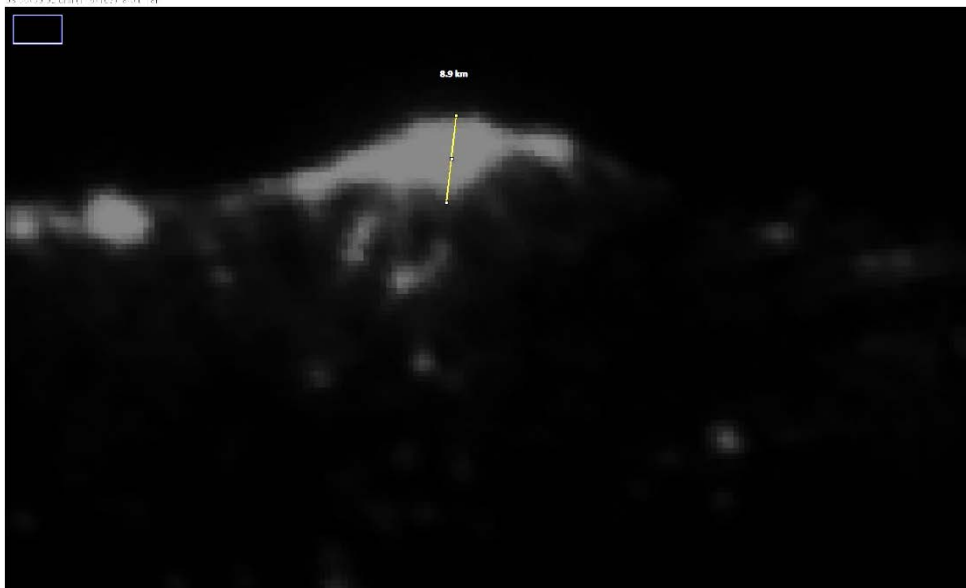
### 8.1 *Urbanisation of the Tripolitania region (Libya)*

Urbanisation in more recent times is well visible in countries that have grown rapidly, both in terms of their economy as well as well as in terms of population. The following two images show the North-West of Libya with the agglomeration of Tripoli in 1992 and 2010. A simple tool is just measuring distances of apparently homogenous areas and to compare the results over time.<sup>17</sup> Measuring the distance of the metropolitan extension between the North (Mediterranean Sea) and the South resulted in some 8.9 km in 1992, while in 2010 the lit area strongly extended up to the desert with 27.56 km.

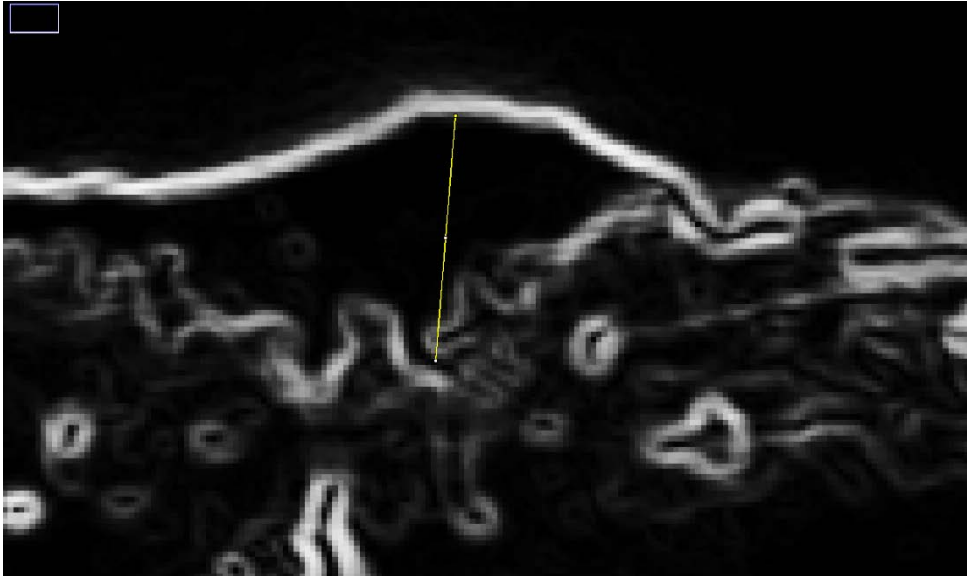
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<sup>16</sup> This is even unusual for untenable areas because of the over-glow effect (cf. the images of Tripoli that contain brighter pixels that are in fact located in the Mediterranean Sea).

<sup>17</sup> ImageJ provides a tool to set scales (here one pixel=0.56 km)



The measurement is, however, imprecise because of over-glow effects. But with ImageJ it is possible to identify edges of areas to better distinguish city light emission from its over-glow into untenable areas, such as the Mediterranean Sea, and thus to correct distance measurements across lit areas (2010 image: 25.2 kilometers).



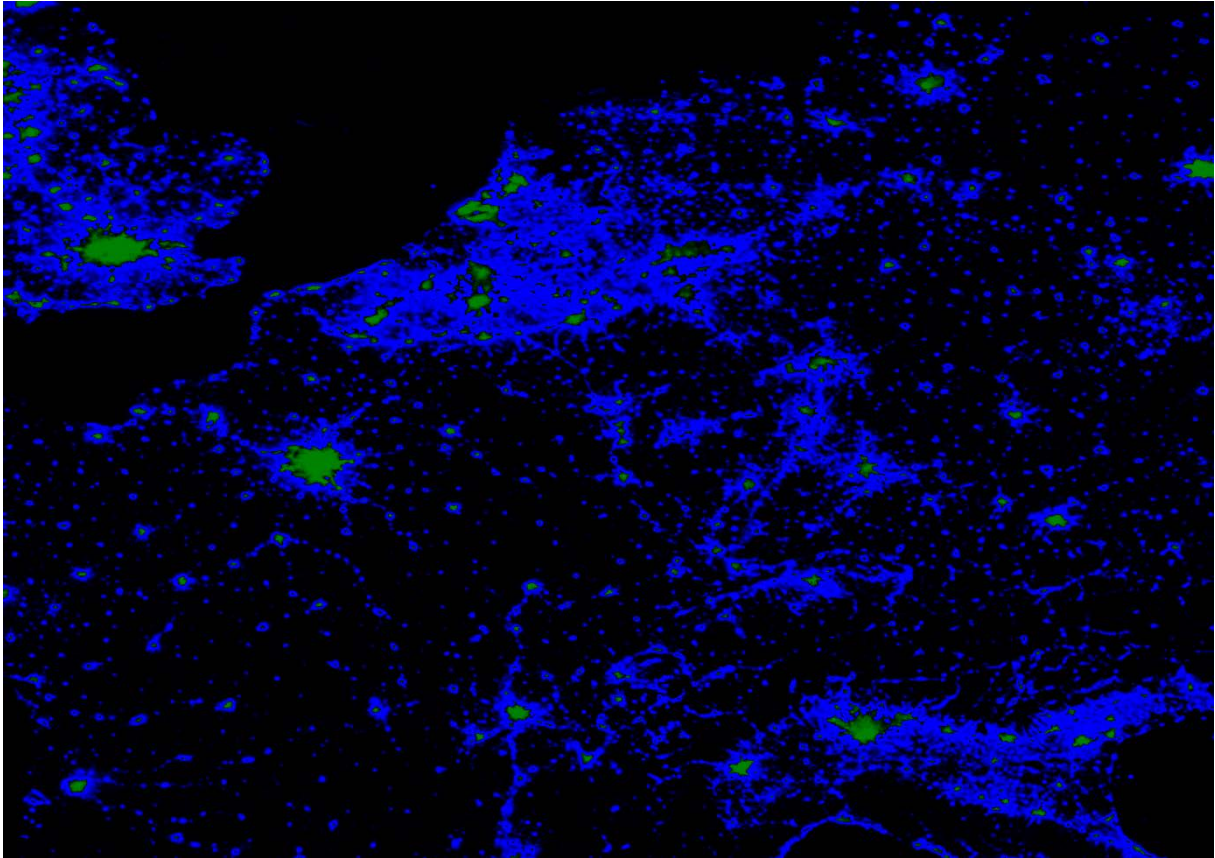
*Tripoli 2010 image: distance across black area between edges is just 25.2 kilometers*

## *8.2 Is there polycentric development? - The Blue Banana*

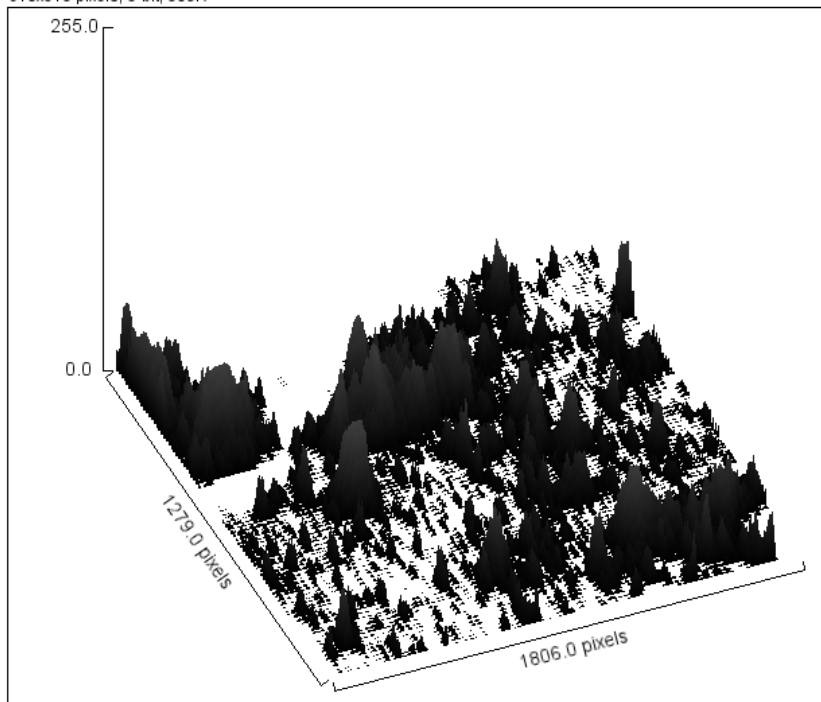
Another simple tool is filtering in making patterns of images more visible. The following images and surface density plots show the Western European core-periphery pattern (Blue Banana) based on luminosity intensity and visualised by the 5-ramps filter. The green areas show those regions with the highest density of top luminosity (DN=63). These are also the major cities and metropolitan regions. The following two images illustrate clearly the map of Western Europe. Both compared show an increasing trend of urbanization (areas of high luminosity have grown) but no polycentric trend, even though a first glimpse could suggest an impression of a significant polycentric trend. The number of clusters with DN > 60 increased from 387 to just 388; by comparing unit areas with DN=63 there is even a reduction from 169 to 155.<sup>18</sup>

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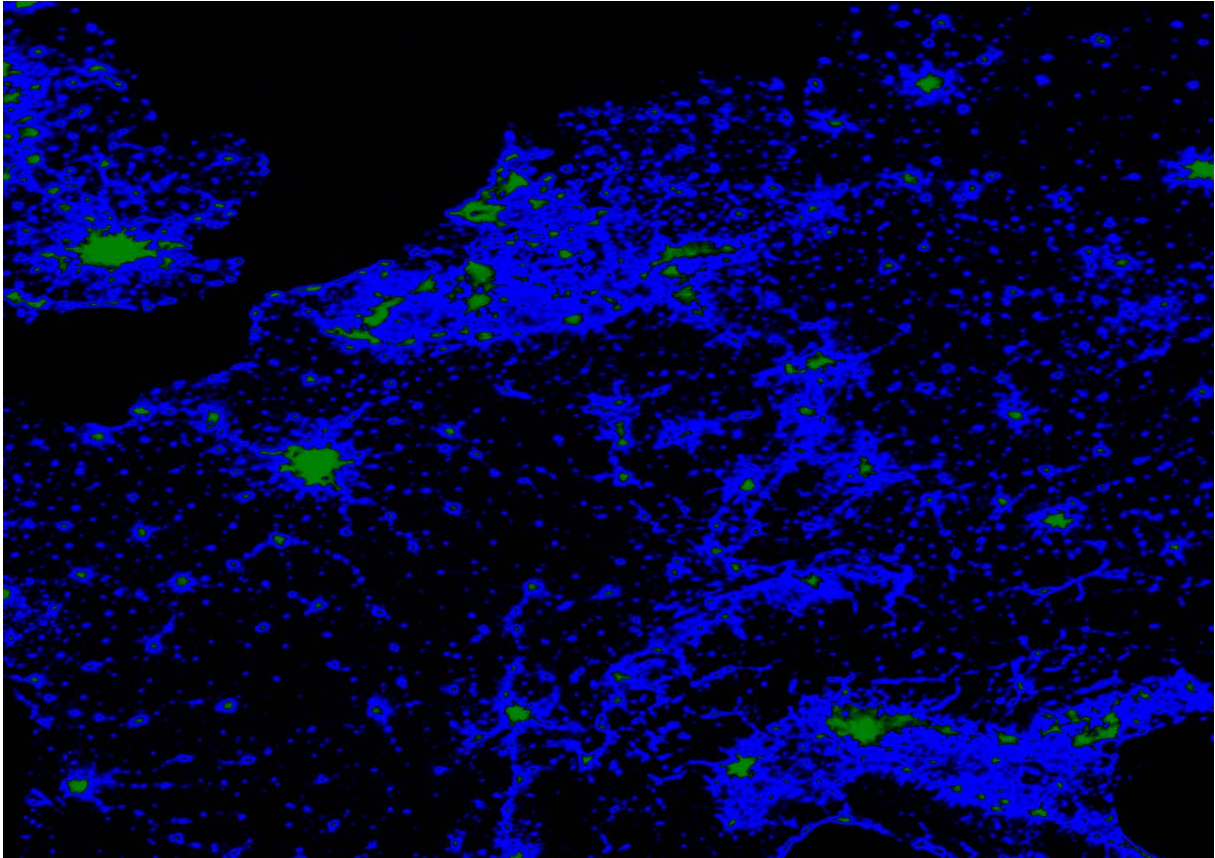
<sup>18</sup> Counting clusters is possible by the "Analyse Particles" command. To run this procedure, the image has first to be transformed into a binary image. Then it is possible to set thresholds, e.g. to select only areas with DN>60.



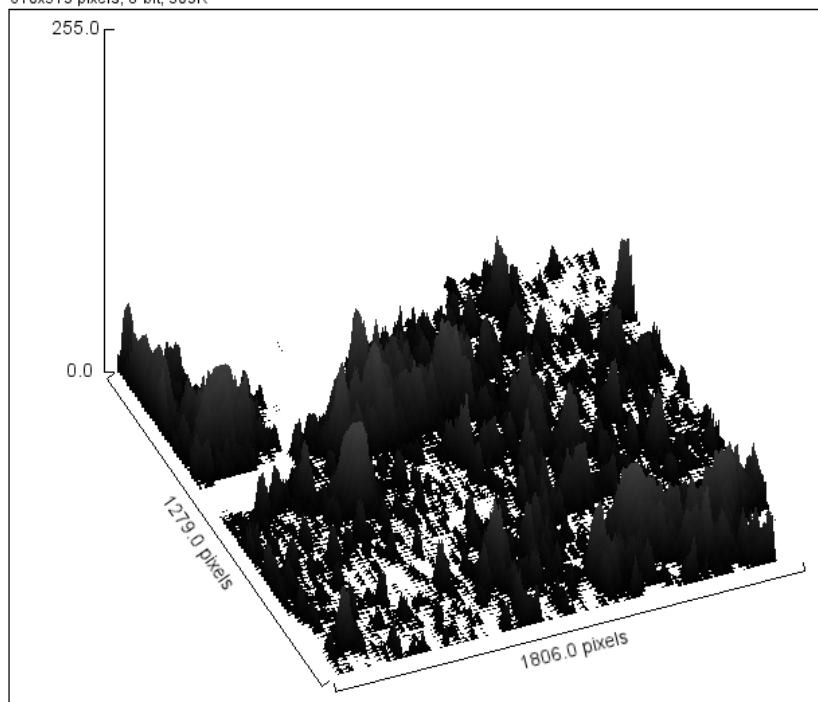
610x519 pixels; 8-bit; 309K



„Blue Banana“ 1992



610x519 pixels; 8-bit; 309K



"Blue Banana" 2009

Beyond pure visual inspection it is of course possible to explore such an image also in statistical terms, such as image correlation, spatial dependence, e.g. *Moran-I*, or spatial heterogeneity, e.g. the rank-size distribution (see further below).

9. *City growth and peri-urbanisation – the meso perspective: Examples Frankfurt and Madrid 1992 versus 2009*

Images contain pixels with different levels of brightness. The DMSP-OLS images are in the range between 0 and 63. Brighter parts are more associated with agglomerations (higher population density); parts with little light emission usually indicate rural areas with lower population density. Areas with zero emission are supposed to be sea, lakes, mountains, forests or other untenable areas. Changing intensity and distribution of light emission may indicate spatial change along the continuum between urban, rural and untenable parts. The observation of different and changing levels of light emission could also help to overcome the traditional normative dichotomy between the “rural” and the “urban”.

9.1 *Frankfurt*

In the following example the region of Frankfurt in 1992 is compared to its shape in 2009. The image shows the wider Rhein-Main area with Frankfurt as its core center. It extends from Gießen in the North to Darmstadt in the South and from Bad Kreuznach in the West to Aschaffenburg in the East. The following descriptive statistics are given by ImageJ:

*Descriptive statistics:*

| Image  | Area   | Mean  | StdDev | Min | Max | IntDen  |
|--|--------|-------|--------|-----|-----|---------|
| Frankfurt1992.v4b.avg_lights_x_pct.lzw-1.tif | 18,832 | 15.15 | 16.86  | 0   | 63  | 285,232 |
| Frankfurt2009.v4b.avg_lights_x_pct.lzw.tif   | 18,832 | 18.63 | 17.45  | 0   | 63  | 350,861 |

The images show exactly the same area at the two points of time (during 17 years). The mean pixel DN and the brightness-weighted average of the x and y coordinates of pixels has increased, the standard deviation has remained rather stable. It is of course to be acknowledged that measured luminosity in the city center is affected by the saturation problem of the non-calibrated images. This may also have an influence on descriptive statistics and the image correlation. The measured correlation of the non-calibrated images is:

*Image correlation 1992 vs. 2009:*

|             |
|-------------|
| Pearson's r |
| 0.96        |

The coefficient is very high indicating either a very little change over that long time period or a very balanced one. The standard deviation remained stable, thus most parts of the area underwent a trend in the same direction. Economic growth has therefore been a spatially balanced process across the area regarded. For comparison, the image correlation of the two images of the Tripoli region (see above) is only at 0.77 indicating a major spatial change over that period.



### *Concentric circle analysis*

Empirical analysis of urbanization has traditionally based on data from official statistics. A core research question has been: How do municipalities develop in terms of population, production and labour market? Sufficient data at municipal level and below are rare, incomplete and often incomparable across municipalities. Analysis of spatial dependence is thus based on data for spatial entities within their administrative boundaries; boundaries and space are irregular. Connectivity weights are constructed on common borders, length of common borders or distances, but the irregular form of the lines of boundaries make it difficult to select a perfect weighting matrix.<sup>19</sup>

An alternative would be to view spatial dependencies across raster units, e.g. square kilometers (see above the raster selection of Frankfurt) or – for micro-spatial neighbourhood analyses – at hectare level, even though over-glow might distort results.

Usually, cities grow radially, only limited by natural or administrative boundaries (like sea, mountains, national borders); areas around cities have various functions (land use), and those functions may change over time. This had been a finding already by von Thunen in the 19<sup>th</sup> century (the *von-Thunen-rings*). For our purpose it could be thus worth to test spatial dependence across such concentric circles. The major advantage is the directly extractable database (pixel intensity distribution) that is independent from administrative borders and adaptable to more natural paths of city/settlement evolution, i.e. a radial spatial change. The selection of the center, the rings, the number of rings and the choice of the radius is at the discretion of the researcher; on the one hand such selection may appear arbitrary, but on the other hand there is the opportunity of an infinite number of different selection choices, so there is the advantage of numerous sensitivity analyses confirming or questioning a prior hypothesis or finding.

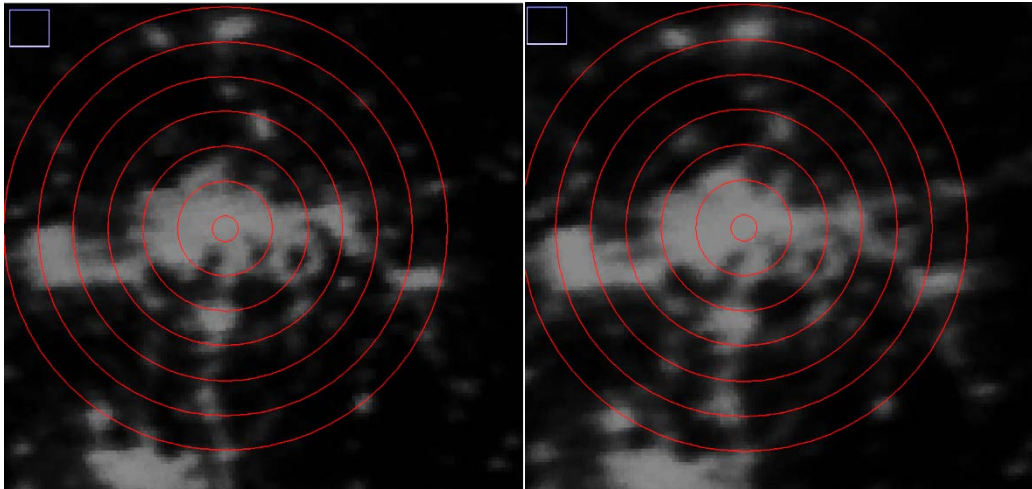
Of course such a concentric circle analysis makes only sense if providing a plausible justification how to position the circles (e.g. city centre + rest of inner city within one versus two circles, separation of outer suburbs and adjacent municipalities, number of circles and decision which circles should be unified to one, different sets for sensitivity analyses etc.).

The following analysis is only a simplified example and its diagnostic value is again limited by the saturation issue of the images:

---

<sup>19</sup> Apart from that, other non-geographical forms of connectivity also influencing spatial dependence are not sufficiently considered (cf. Beck et al 2006).

*Night Satellite Images: Frankfurt 1992 and 2009:*



*Frankfurt 1992*

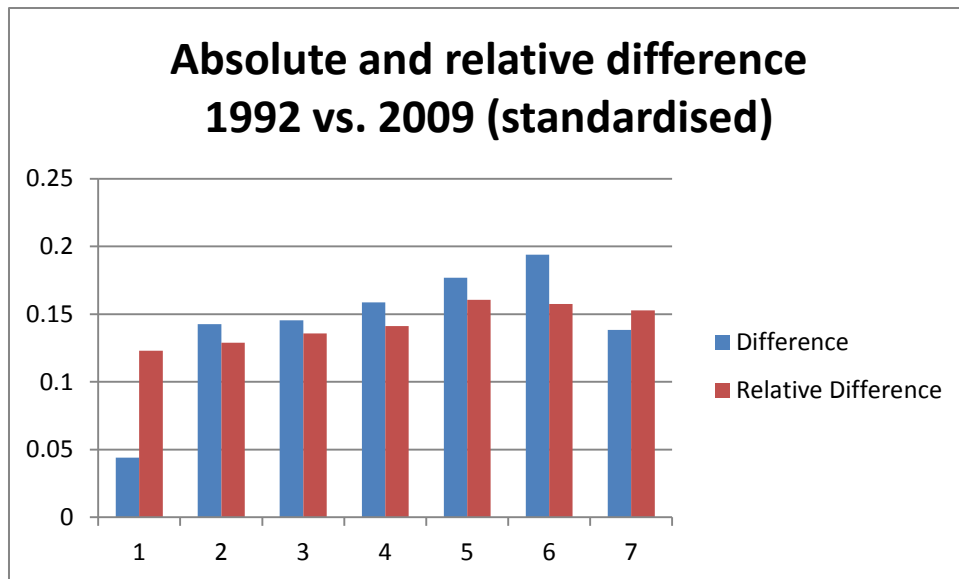
*Frankfurt 2009*

*Descriptive statistics of the concentric circle analysis 1992:*

| 1992 | Radius | Samples | Mean 1992 |
|------|--------|---------|-----------|
| 1    | 3.8    | 33      | 59.93939  |
| 2    | 14     | 113     | 49.0531   |
| 3    | 24.2   | 197     | 26.8934   |
| 4    | 34.4   | 277     | 21.44404  |
| 5    | 44.6   | 361     | 12.27701  |
| 6    | 54.8   | 441     | 14.53515  |
| 7    | 65     | 521     | 11.94242  |

*Comparison 2009 versus 1992:*

| Mean 2009 | Difference | Mean diff. % |
|-----------|------------|--------------|
| 60.939    | 1          | 102          |
| 52.292    | 3.239      | 107          |
| 30.193    | 3.299      | 112          |
| 25.043    | 3.599      | 117          |
| 16.294    | 4.017      | 133          |
| 18.937    | 4.401      | 130          |
| 15.084    | 3.142      | 126          |



Absolute differences are smallest for the inner circle. This might be explained by the saturation of lights with highest luminosity and the smaller number of observations in this area. Relative differences are rather balanced, being larger in the outer rings (5,6, and 7), showing that some detectable change takes more place in the peri-urban and rural contiguous space around Frankfurt.

The 2<sup>nd</sup> and 3<sup>rd</sup> circle represents major parts of the peri-urban belt, while the 5<sup>th</sup> to 7<sup>th</sup> circle cover mostly rural districts in the contiguous territory. Hence, rural areas have further integrated in the peri-urban landscape of Frankfurt and eventually the Rhein-Main area

There is one important caveat in this interpretation. Even though the Rhein-Main area is dominated by its urban center Frankfurt, the region is nevertheless rather characterized by a polycentric urban distribution with other major cities such as Wiesbaden, Mainz and Darmstadt. Hence, this agglomeration trend is not alone to be attributed to the pull-factor of Frankfurt.

A further *Moran's local I* analysis (LISA)<sup>20</sup> could shed more light on the changes of the local autocorrelation along the rings around the city. However, the results on such a small number of observations are not sufficiently reliable<sup>21</sup>. For an empirical analysis with *Moran's local I* one would need to increase the number of rings (minimum 30) to identify the geographical "breaks" of the coefficients along the rings. A shift of such breaks over time will indicate how areas have changed their functional relations. Further below, a comparative LISA analysis for the Madrid agglomeration is illustrated; this region has exhibited stronger spatial change than the Frankfurt region.

<sup>20</sup> A positive *Moran's local I* indicates an area with neighbours displaying similar value levels of a variable. A negative coefficient represents an outlier. Here the variable values are highly dissimilar between two neighbouring areas. *Moran's local I* can only be properly interpreted under consideration of its z-score.

<sup>21</sup> Because of the little number of areas there is no meaningful significance test, so that here the Moran I coefficient can be only regarded as a descriptive correlation measurement. (Cf. Schulze 1995)

## 9.2 Madrid

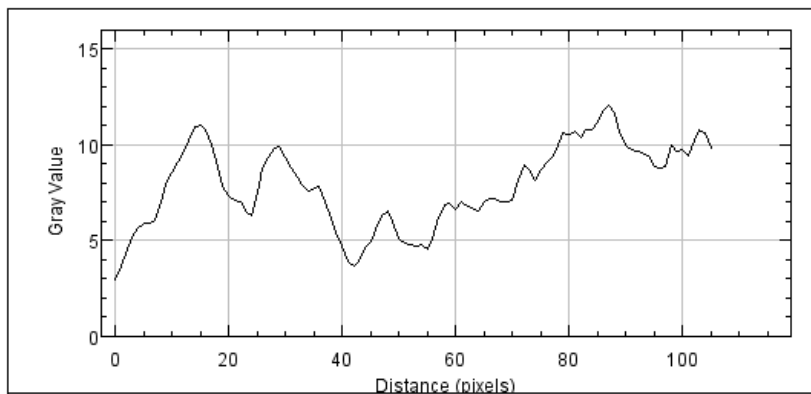
A second example selected is the agglomeration of Madrid. This agglomeration with around 7 million inhabitants is significantly larger than that of Frankfurt.

Descriptive statistics suggest a slightly more dynamic growth than that of Frankfurt. Mean pixel-DN increased from a similar level of 16.61 to 24.19 (e.g. a change of 45% versus 23% in the case of Frankfurt). Image correlation is thus also lower than in the case of Frankfurt.

*Descriptive statistics:*

| Image   | Area   | Mean  | StdDev | Min | Max | IntDen  |
|---|--------|-------|--------|-----|-----|---------|
| Madrid-F101992.v4b.avg_lights_x_pct.lzw-1.tif | 18,832 | 16.61 | 19.47  | 0   | 63  | 312,715 |
| MadridF162009.v4b.avg_lights_x_pct.lzw.tif    | 18,832 | 24.19 | 21.50  | 0   | 63  | 455,588 |

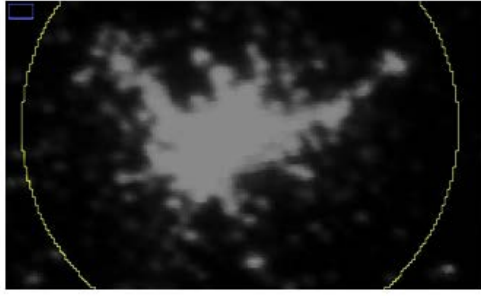
*Difference plot profile 1992 and 2009 (mean difference per pixel column):*



*Image correlation 1992 vs. 2009:*

|             |
|-------------|
| Pearson's r |
| 0.91        |

This is further examined by the radial profile analysis, another ImageJ plugin. The following two snapshots show the same section of Madrid agglomeration in 1992 and 2009. By visual inspection alone it becomes obvious that there is a considerably higher density of lights in 2009, revealing a process of urban growth.

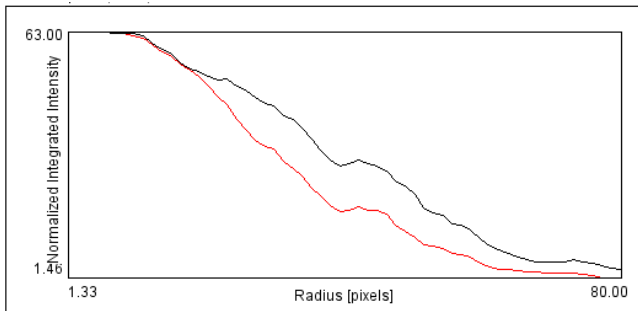


*Madrid agglomeration 1992*



*Madrid agglomeration 2009*

The slope of the pixel distribution within the radial profile has significantly changed during 1992-2009 as shown by the following figures.<sup>22</sup>



*Madrid agglomeration: Radial profile  
1992 (red) and 2009 (black)*

```

Madrid 1992:
Sum of residuals squared: 1547.59839
Standard deviation: 5.12157
R2: 0.94784
OLS: y = 64.87989 - 0.93758x

Madrid 2009:
Sum of residuals squared: 360.71150
Standard deviation: 2.47260
R2: 0.98618
OLS: y = 69.02228 - 0.89700x

```

The above comparison of the same areal section shows a stronger trend to a linear relationship of radius and pixel density in 2009, while in 1992 the relation was more convex<sup>23</sup>

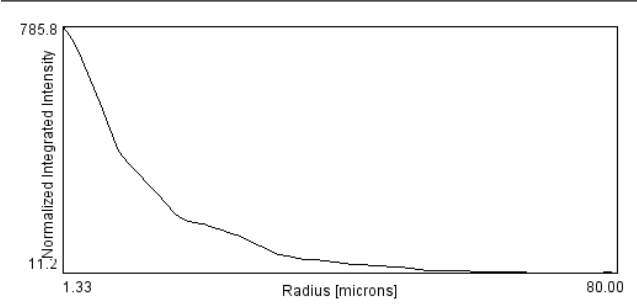
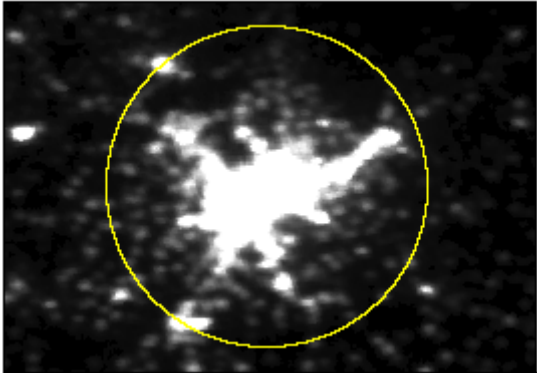
<sup>22</sup> Normalised integrated intensity (Y): normalized by the average light intensity per pixel to sense pixel-to-pixel relative homogeneity

<sup>23</sup> showing the shape of  $f(x) < 0$  and  $f'(x) > 0$

revealing a decreasing marginal appearance of darker items (rural) along the radius from the inner to the outer sections. The 2009 image shows a slightly more balanced distribution of the pixel DN along the spectrum between 63 in the center and 0 in the outer areas, confirming that the agglomeration in the center has grown; therefore the linear best-fit line is better matched than that for the 1992 image. Consequently there was also a slightly weaker goodness of fit of the linear regression for 1992 ( $R^2=0.95$ ) compared to that of 2009 ( $R^2=0.99$ ).<sup>24</sup>

In empirically selecting a threshold between urban and rural luminosity (by randomly selecting rural and urban zones of a country) it is possible to observe the transition points of urbanizing rural zones. E.g. if selecting the threshold for Spain at, let's say, DN=10, areas in the distance of 40 kilometers (67 pixels) around the centre of Madrid were typically rural in 1992 but have become peri-urban (on average) in 2009.

When using radiance-calibrated images, the shape, the coefficient and the fit are much different, shedding more light on the distribution of lights and range of light intensity in the core city. The typical *Pareto* distribution of light emission at global level then becomes also visible for metropolitan areas.



*Madrid agglomeration 2011: Radial profile (radiance-calibrated image)*

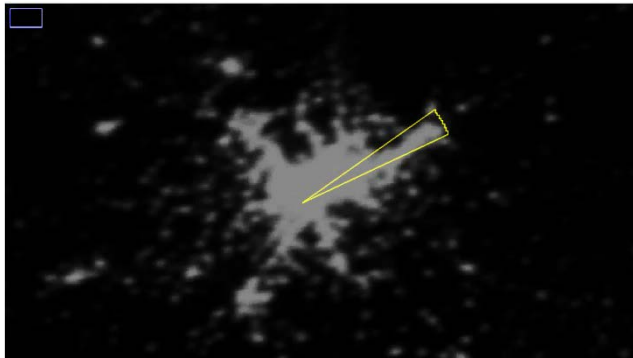
```

Madrid 2011 (radiance-calibrated image):

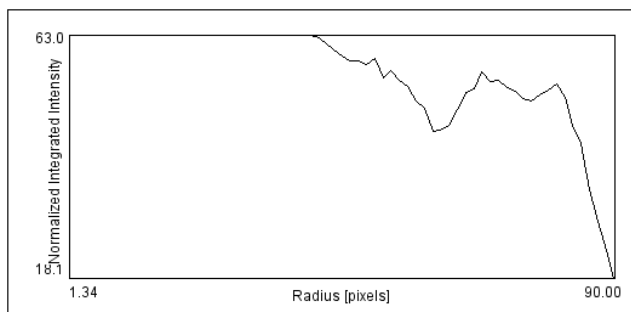
Sum of residuals squared: 840659.269
Standard deviation: 119.36692
R2: 0.62958
Parameters (OLS): y = a + bx
    a = 411.71501
    b = -6.68308
    
```

<sup>24</sup> For a better fit of convex or concave curves, the numerical list of values plotted along the radius can be used to find a polynomial approximation of shape.

With increasing distance from the city center, the number of observations (pixels) and the spread of pixel-DN increases; mean values are increasingly disturbed by outliers. In analogy to Hoyt's (1941) urban sector theory it is then possible to analyse the radial profile within urban segments and to subdivide the circle around the city into "pieces of cake", for instance highlighting development axes of major interest, such as the example below:



*Madrid agglomeration 2009: Radial profile of the Madrid North-East axis*



Madrid 2009:

```

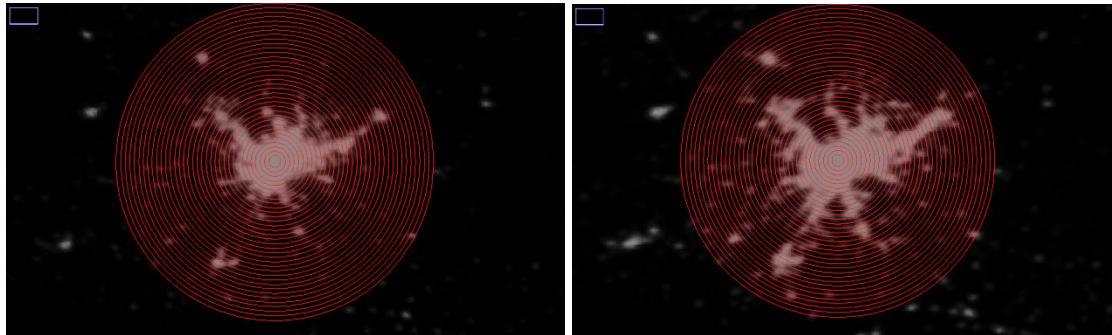
Sum of residuals squared: 2355.81979
Standard deviation: 5.97447
R2: 0.61524
Parameters (OLS): y = a + bx
      a = 69.32640
      b = -0.28864
  
```

In this case, we find a completely different shape of the slope with nearly half of the pixels top-coded. It is possible to better detect polycentric structures (and their change) along different axes of urbanization. While the full circle analysis shows an almost monotonously declining curve, the partial analysis includes significant breaks with increasing light emission along the radius. A static comparison of such breaks in different selected sectors at two or more points of time may shed light on the specific change of spatial axes of urbanization and/or polycentric evolution. Such analyses are possibly useful to complement other empirical studies on urban spatial change. For that purpose, the selection of sectors, the size of the radius and the respective angle need to be well specified.

#### *Spatial autocorrelation: Local Moran-I Analysis*

In a further analysis I looked whether the city and its periphery has become more integrated during the 1992-2009 period and tested it with a spatial autocorrelation analysis across concentric circles, going beyond that demonstrated for Frankfurt. For the Madrid case, a

selection of 35 concentric circles is specified, looking at the global autocorrelation coefficient (Moran's I) and, particularly, at the local coefficients (LISA).<sup>25</sup> It was first tested with a simple binary weight matrix, i.e. weighting contiguous circles with one and all others with zero. However a non-binary distance matrix resulted in more significant coefficients, which better illustrate the changes that took place between 1992 and 2009.



Madrid 1992: 35 circles

Madrid 2009: 35 circles

| Moran's Ii (Mean)Madrid 1992 (non-binary weight matrix) |              |              |              |              |
|---|--------------|--------------|--------------|--------------|
| Location  | Ii           | sd(Ii)       | z            | p-value*     |
| 13  | 0.002        | 0.498        | 0.062        | 0.475        |
| 14  | 0.003        | 0.498        | 0.065        | 0.474        |
| 15  | 0.024        | 0.498        | 0.108        | 0.457        |
| 16  | 0.056        | 0.498        | 0.171        | 0.432        |
| 12  | 0.062        | 0.498        | 0.184        | 0.427        |
| 17  | 0.119        | 0.498        | 0.299        | 0.383        |
| 11  | 0.157        | 0.498        | 0.374        | 0.354        |
| 18  | 0.203        | 0.498        | 0.466        | 0.321        |
| 19  | 0.286        | 0.498        | 0.634        | 0.263        |
| 10  | 0.334        | 0.498        | 0.731        | 0.232        |
| 20  | 0.370        | 0.498        | 0.802        | 0.211        |
| 21  | 0.475        | 0.498        | 1.014        | 0.155        |
| 22  | 0.546        | 0.498        | 1.156        | 0.124        |
| 24  | 0.552        | 0.498        | 1.168        | 0.121        |
| 23  | 0.561        | 0.498        | 1.186        | 0.118        |
| 25  | 0.563        | 0.498        | 1.191        | 0.117        |
| 26  | 0.606        | 0.498        | 1.278        | 0.101        |
| 35  | 0.906        | 0.724        | 1.293        | 0.098        |
| 9   | 0.644        | 0.498        | 1.352        | 0.088        |
| 27  | 0.655        | 0.498        | 1.376        | 0.084        |
| 28  | 0.706        | 0.498        | 1.478        | 0.070        |
| 29  | 0.764        | 0.498        | 1.595        | 0.055        |
| 34  | 0.904        | 0.574        | 1.626        | 0.052        |
| 30  | <b>0.809</b> | <b>0.498</b> | <b>1.685</b> | <b>0.046</b> |
| 31  | <b>0.854</b> | <b>0.498</b> | <b>1.776</b> | <b>0.038</b> |
| 32  | <b>0.877</b> | <b>0.498</b> | <b>1.820</b> | <b>0.034</b> |
| 33  | <b>0.899</b> | <b>0.498</b> | <b>1.865</b> | <b>0.031</b> |
| 8   | <b>1.082</b> | <b>0.498</b> | <b>2.234</b> | <b>0.013</b> |
| 7   | <b>1.668</b> | <b>0.498</b> | <b>3.410</b> | <b>0.000</b> |
| 6   | <b>2.292</b> | <b>0.498</b> | <b>4.665</b> | <b>0.000</b> |
| 1   | <b>3.526</b> | <b>0.724</b> | <b>4.913</b> | <b>0.000</b> |
| 5   | <b>2.808</b> | <b>0.498</b> | <b>5.700</b> | <b>0.000</b> |
| 2   | <b>3.510</b> | <b>0.574</b> | <b>6.166</b> | <b>0.000</b> |
| 4   | <b>3.233</b> | <b>0.498</b> | <b>6.554</b> | <b>0.000</b> |
| 3   | <b>3.439</b> | <b>0.498</b> | <b>6.970</b> | <b>0.000</b> |

-----

E(Ii) = -0.029  
 Moran's I = 0.986  
 \*1-tail test; significant coefficients in bold numbers

<sup>25</sup> The selection of size and number of circles is at the discretion of the researcher. In ImageJ the circles have the same radial distance. If selecting a small number of circles, there may be not only the problem of a too little number of observations but also misspecification, because areas of functional roles are not grouped around a city in constant distances (e.g. the von Thunen or Burgess rings). However, when increasing the number of circles it becomes possible to compare different radial sizes of the rings, simply by adding up the moments of certain contiguous circles and calculating the mean pixel-DN. Then again it is possible to estimate spillovers by spatial autocorrelation.



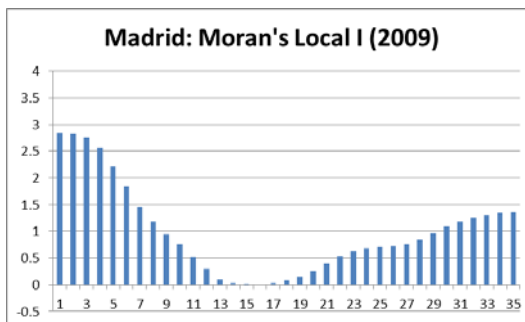
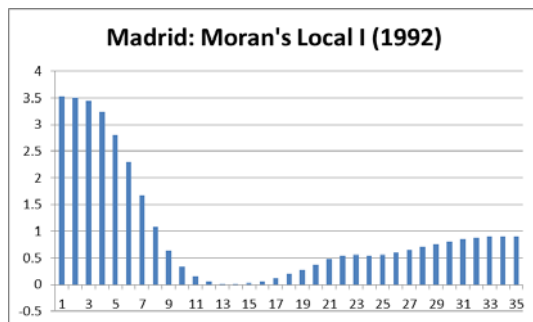
Moran's Ii (Mean): Madrid 2009 (non-binary weight matrix)

| Location  | Ii           | sd(Ii)       | z            | p-value*     |
|-----------|--------------|--------------|--------------|--------------|
| 16        | -0.002       | 0.501        | 0.055        | 0.478        |
| 15        | 0.008        | 0.501        | 0.076        | 0.470        |
| 17        | 0.023        | 0.501        | 0.104        | 0.459        |
| 14        | 0.024        | 0.501        | 0.106        | 0.458        |
| 18        | 0.082        | 0.501        | 0.222        | 0.412        |
| 13        | 0.090        | 0.501        | 0.238        | 0.406        |
| 19        | 0.154        | 0.501        | 0.367        | 0.357        |
| 20        | 0.251        | 0.501        | 0.560        | 0.288        |
| 12        | 0.298        | 0.501        | 0.654        | 0.257        |
| 21        | 0.397        | 0.501        | 0.852        | 0.197        |
| 11        | 0.526        | 0.501        | 1.109        | 0.134        |
| 22        | 0.531        | 0.501        | 1.118        | 0.132        |
| 23        | 0.622        | 0.501        | 1.301        | 0.097        |
| 24        | 0.681        | 0.501        | 1.418        | 0.078        |
| 25        | 0.706        | 0.501        | 1.468        | 0.071        |
| 26        | 0.714        | 0.501        | 1.485        | 0.069        |
| 10        | 0.752        | 0.501        | 1.560        | 0.059        |
| 27        | 0.753        | 0.501        | 1.562        | 0.059        |
| <b>28</b> | <b>0.840</b> | <b>0.501</b> | <b>1.736</b> | <b>0.041</b> |
| <b>35</b> | <b>1.355</b> | <b>0.729</b> | <b>1.900</b> | <b>0.029</b> |
| <b>9</b>  | <b>0.944</b> | <b>0.501</b> | <b>1.944</b> | <b>0.026</b> |
| <b>29</b> | <b>0.964</b> | <b>0.501</b> | <b>1.983</b> | <b>0.024</b> |
| <b>30</b> | <b>1.084</b> | <b>0.501</b> | <b>2.223</b> | <b>0.013</b> |
| <b>34</b> | <b>1.338</b> | <b>0.578</b> | <b>2.368</b> | <b>0.009</b> |
| <b>31</b> | <b>1.177</b> | <b>0.501</b> | <b>2.409</b> | <b>0.008</b> |
| <b>8</b>  | <b>1.188</b> | <b>0.501</b> | <b>2.431</b> | <b>0.008</b> |
| <b>32</b> | <b>1.249</b> | <b>0.501</b> | <b>2.552</b> | <b>0.005</b> |
| <b>33</b> | <b>1.304</b> | <b>0.501</b> | <b>2.663</b> | <b>0.004</b> |
| <b>7</b>  | <b>1.467</b> | <b>0.501</b> | <b>2.987</b> | <b>0.001</b> |
| <b>6</b>  | <b>1.845</b> | <b>0.501</b> | <b>3.743</b> | <b>0.000</b> |
| <b>1</b>  | <b>2.843</b> | <b>0.729</b> | <b>3.942</b> | <b>0.000</b> |
| <b>5</b>  | <b>2.221</b> | <b>0.501</b> | <b>4.494</b> | <b>0.000</b> |
| <b>2</b>  | <b>2.824</b> | <b>0.578</b> | <b>4.939</b> | <b>0.000</b> |
| <b>4</b>  | <b>2.563</b> | <b>0.501</b> | <b>5.176</b> | <b>0.000</b> |
| <b>3</b>  | <b>2.756</b> | <b>0.501</b> | <b>5.562</b> | <b>0.000</b> |

E(Ii) = -0.029

Moran's I = 0.988

\*1-tail test; significant coefficients in bold numbers



In both years there is a strong and highly significant local autocorrelation for the most inner circles and moderately significant ones for most outer circles. The above results show that there are more significant<sup>26</sup> coefficients in 2009 compared to 1992. In 1992 there were twelve circles with significant coefficients, while the number increased to 17 in 2009. As the above diagrams show, the level of spatial autocorrelation decreased for the first inner circles, but increased for circles 8 and 9 and the outer periphery<sup>27</sup>. A typical indication of spatial integration around the city of Madrid is the increase of the *local Moran's I* for circle 9 from

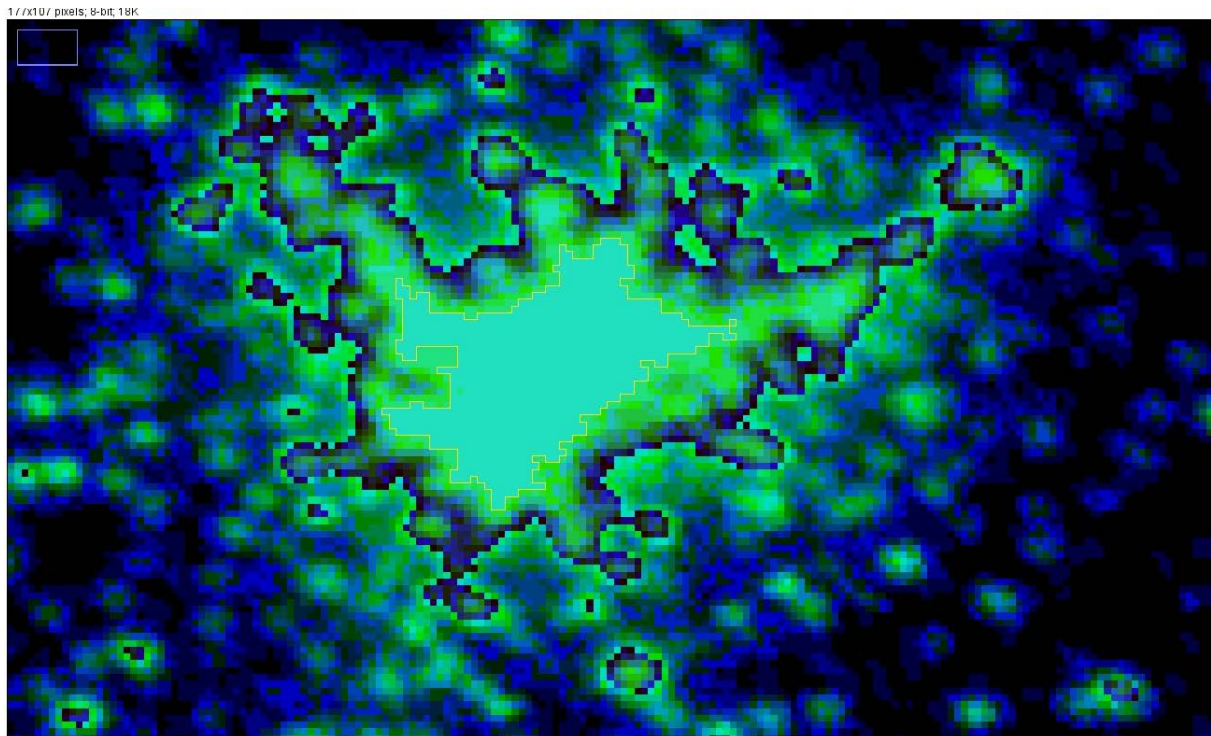
<sup>26</sup> By setting the threshold at  $p=0.05$

<sup>27</sup> The local coefficients are directly related to Moran's global I via their average. Hence, as global Moran's I has remained constant, the bars in the diagrams above can be somehow interpreted as "communicating pipes".

insignificant 0.644 in 1992 to significant 0.944 in 2009. While in 1992 coefficients for the two outer rings were insignificant, they became significant by 2009. Local *Moran-I* coefficients of the middle circles are insignificant. They represent areas close to the average luminosity of the selected agglomeration. However, the shift of insignificant coefficients to the right indicates a net centrifugal force (the minimum shifted from circle 13 to circle 16). For 2009 there is an increase of significant coefficients at the margins of the inner third of rings, suggesting that areas in that distance to the central city become functionally more integrated, also confirmed by the shapes of the radial profile plots above<sup>28</sup>. Hence, the reduction of the number of insignificant coefficients reveals that outer areas of the agglomeration are gradually coalescing with the inner ones.

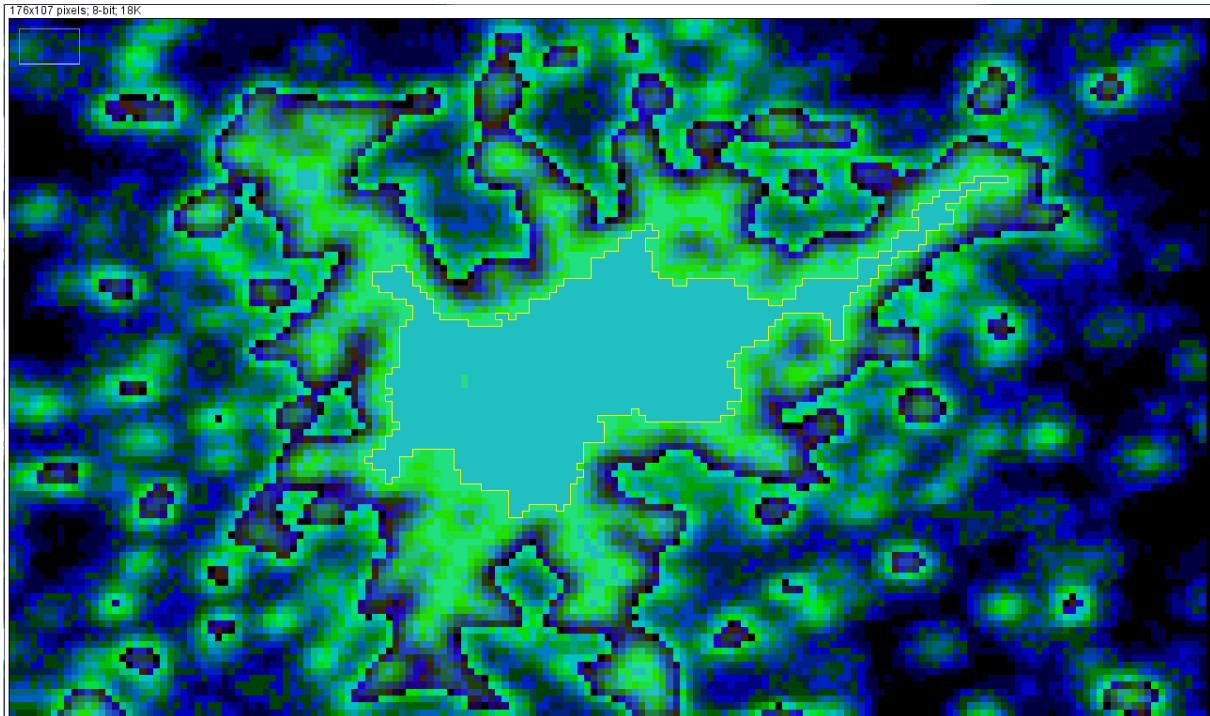
#### *Further tools for the analysis of urban spatial change*

A last simple but interesting tool of ImageJ, discussed here, is the Wand tool. With that instrument it is e.g. possible to view changes of the inner city or specific parts of the city by capturing the areas with identical pixel intensity – in this case the range of top-coded values ( $DN \geq 63$ ), assuming that this level represents luminosity of a city centre. By using the 3-3-2 RGB filter and subsequently using the Wand tool, the boundaries of the inner city (top-coded) and the wider city are made visible.<sup>29</sup> Moments of pixels can be calculated. The results show that the city area has grown strongly over the time, especially integrating the agglomeration along the Highway E-90 North-East (Torrejón de Ardoz to Guadalajara).



<sup>28</sup> Theoretically *Moran's I* could also increase if there is a negative trend of luminosity in all sections, with other words, when cities shrink or "ruralise". A decrease of *Moran's I* would happen if there are contrary trends in contiguous sections.

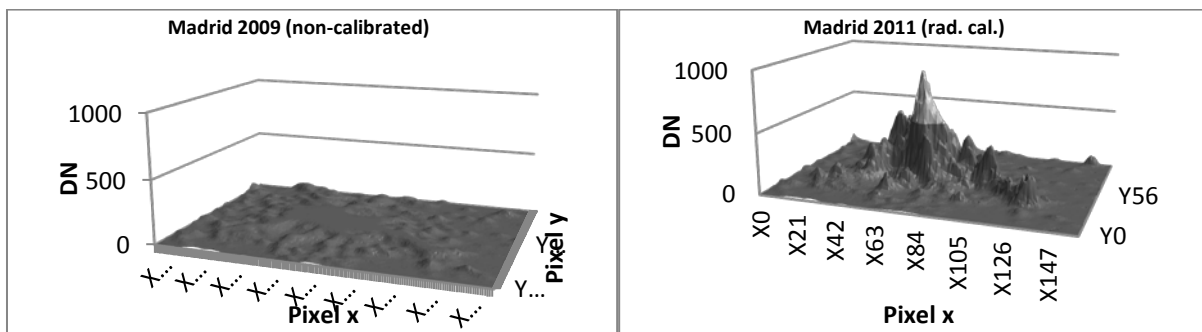
<sup>29</sup> The 3-3-2 RGB filter is capable to show edges, i.e. fault lines, beyond those pixel intensity declines. With ImageJ it is also possible to directly identify edges and maxima within their segments (see earlier: the image of Tripoli 2010).



*Spatial growth of Madrid city 1992 vs. 2009*

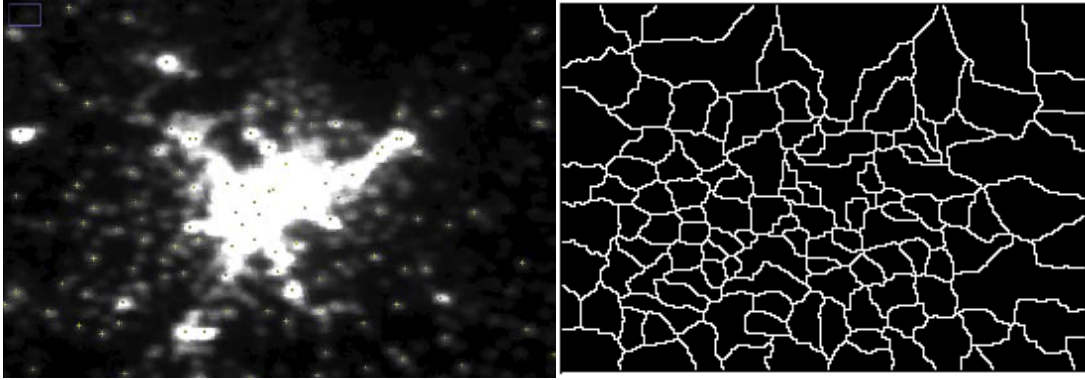
| Label                                       | Area | Mean     | StdDev  | Mode | Min | Max |
|---|------|----------|---------|------|-----|-----|
| Madrid-F101992.v4b.avg_lights_x_pct.lzw.tif | 866  | 62.99885 | 0.03398 | 63   | 62  | 63  |
| MadridF162009.v4b.avg_lights_x_pct.lzw.tif  | 1411 | 62.99858 | 0.03764 | 63   | 62  | 63  |

Within the area of top-coded DN values the real distribution is highly different from that measured in the non-calibrated images. The following comparison of density plots between a non-calibrated and a radiance-calibrated image gives some idea on that:



*Density plots of non-calibrated and radiance-calibrated images compared*

For the calibrated version, local maxima of light emission (here within a tolerance of 10) and their segmentation can then more precisely shed light on the distribution of local radiance activity.



There are further analytic image analysis tools that could be tested for spatial analysis. Worth mentioning are fractal box counts, nearest neighbour analysis and others. In the final part of this paper I like to highlight a feature that is related to the deeper underlying order of spatial heterogeneity, namely the rank-size distribution of cities (or settlements respectively).

#### 10. Spatial heterogeneity and Zipf's law

Further to the analysis of spatial dependence and the variation, scanning of night imagery can also be applied to analyses of spatial heterogeneity and to detect regular patterns of distribution. A well-known quasi-natural law of human settlement is the so-called rank-size distribution of cities. *Zipf's law* (Zipf 1949)<sup>30</sup> says that, at any country level, the largest city is  $x$  times larger than the city on rank  $x$  (measured on a logarithmic scale). Hence, based on empirical observations, the rank size distribution of cities is universally:

$$\log(S) = \log(c) - q \log(r) + u$$

where  $S$  means size of the city,  $c$  is a constant,  $r$  is the rank and the coefficient

$$q \approx 1$$

With  $q=1$  the reverse relationship (*Pareto* distribution)

$$r = cS^{-q}$$

would be equal, so it is possible to take both,  $S$  as well as  $r$  as the dependent variable or predictor respectively.

There is a large volume of research on *Zipf's law*, but this is overwhelmingly based on official data that is highly distorted by administrative boundaries, changes and error in official statistics. Such artificial boundaries – especially if they change – lead to major distortions of naturally delineated settlements. Thus there are also major deviations of validity of *Zipf's law*

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<sup>30</sup> *Zipf's law* is actually based on observations of Felix Auerbach in 1913 about patterns of population distribution. There are others before Zipf, such as Alfred Lotka or Frank Benford who have found similar relationships in other areas of life where observations are *pareto*-distributed.

across countries and during time. Many observers thus have questioned the validity of *Zipf's* law.

Jiang et al. (2015) have used night satellite imagery to test *Zipf's* law globally. Luminosity is thus used as a proxy for settlements. The important advantage is that, with the data of light emission, there are neither administrative data restrictions nor do we need to acknowledge administrative boundaries. There is worldwide access to this variable. Furthermore it is possible to delineate functional settlements that make up “natural cities”. A crucial point in this connection is the property of the *Pareto*-distribution, namely that there are many more small items than big ones and that the distribution of observations below and above the mean is rather skewed so that the mean, as an expected value of a sample, becomes meaningless.<sup>31</sup> However, the mean can be used to identify thresholds. It is possible to view the distribution of pixel intensity (DN) of a country or region and to find the mean DN. In subdividing the two parts (below and above the mean) into heads and tails and repeating this procedure for the observations for the respective head sections, the distribution becomes more and more linear, so that the procedure converges to a 50:50 distribution of observations below and above the mean. Settlements identified in this linear distribution are considered natural cities. Filtering-out pixel-DN being lower than their respective mean will then lead to a reduction of settlements that qualify as cities. Interestingly, the pattern of distribution of natural cities is significantly in line with *Zipf's* law, and this not only at country level but also at global and continental level. (Jiang 2015).

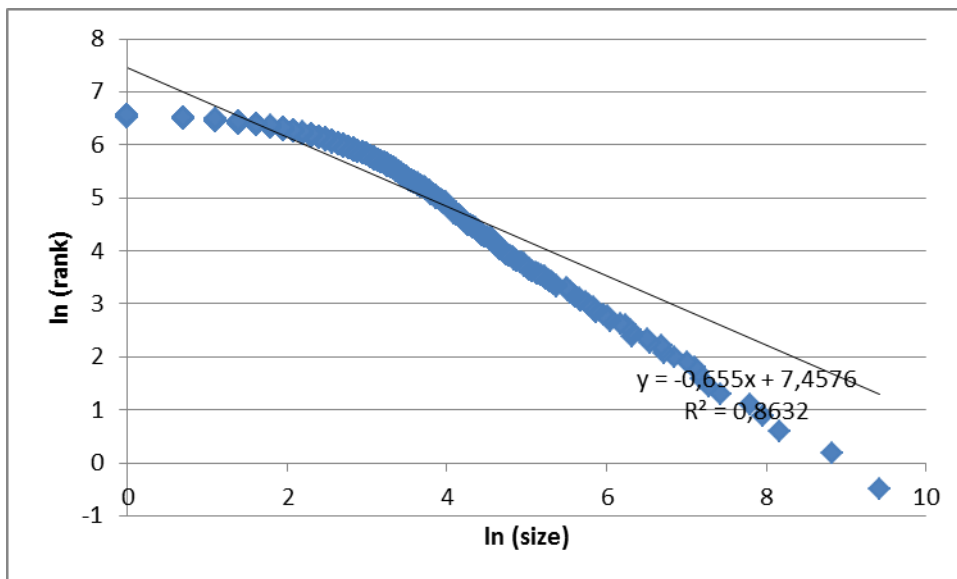
I tested this approach for Germany 2009 (within the area of 47.5° to 55.5° latitude and 6.5° to 14.5° longitude) with ImageJ

| Round     | Pixels  | Mean DN | Obs. Above mean | Obs. Below mean |
|-----------|---------|---------|-----------------|-----------------|
| 1 (total) | 689,984 | 11.55   | 32.3            | 67.7            |
| 2 (head)  | 222,869 | 27.76   | 41.6            | 58.4            |

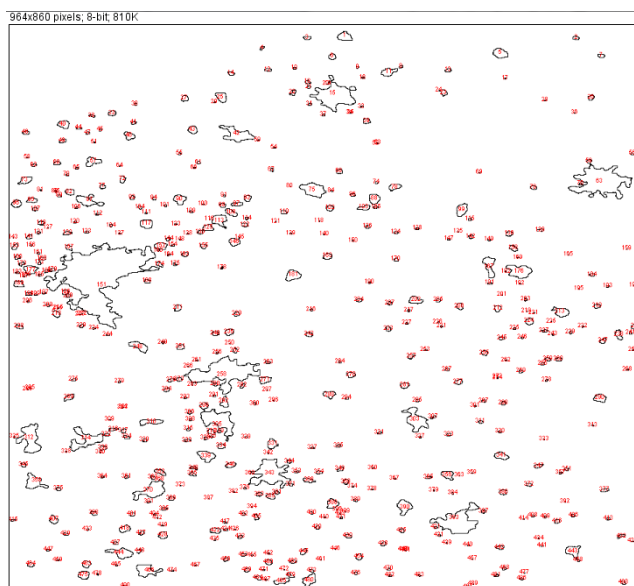
First a *Zipf* regression on the total 689,984 pixels with DN>0 was executed. The estimation of  $q$  was done with the *Gabaix-Ibragimov*-estimator<sup>32</sup> for the *Zipf* regression to reduce the bias in the standard OLS. Therefore rank is regressed on size instead size on rank. As we see, by viewing the total number of observations, the estimate for  $q$  differs strongly from the expected -1:

<sup>31</sup> For distributions with big tails (such as *Pareto*-, log-normal- or *Levy*- distributions) the central limit theorem does not hold.

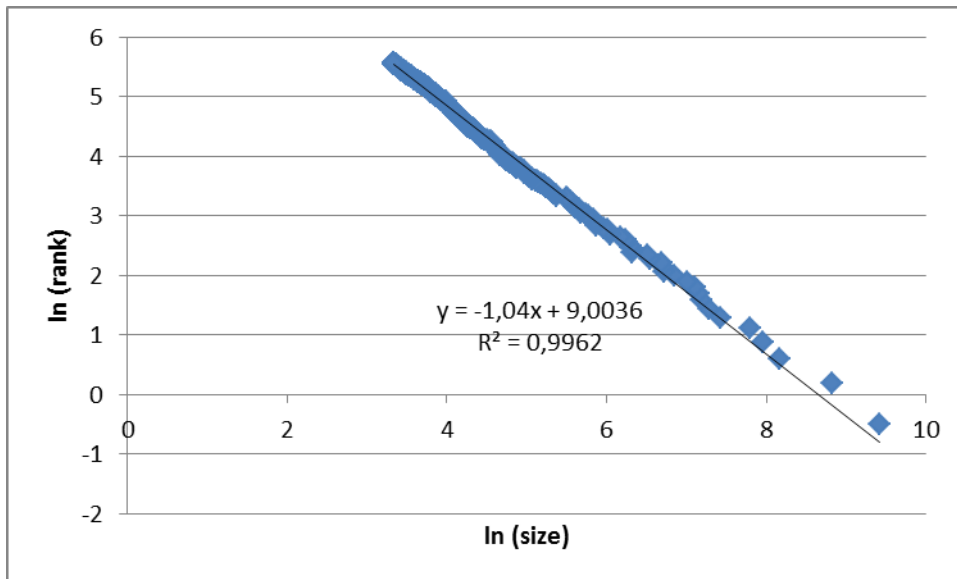
<sup>32</sup> This coefficient is obtained by reducing the rank by  $\frac{1}{2}$  and then using the standard OLS approach, i.e.  $\ln(r-1/2) = c-q \ln(S)$ . The standard error of the Pareto exponent differs from the OLS standard error. It is asymptotically  $(2/n)^{1/2}q$  (cf. Gabaix and Ibragimov 2007).



Following Jiang's approach to select a "natural" delineation of settlements any pixels lower than the threshold of DN 28 were filtered out. This generated the following map by ImageJ:



The subsequent *Zipf* regression led to the following result, which is significantly closer to the expected -1, with a nearly perfect goodness of fit:



With a regression coefficient of -1.04, *Zipf's* law can be confirmed for the selected area.

Using the tool of night imagery it could be interesting to undertake further research on the behavior of the rank-size distribution of settlements among numerous different sizes and forms of selection of areas - within a country or for transnational regions, and to analyse the variation of *Zipf* coefficients among different shapes and sizes of areas selected.

### 11. Conclusion

The DMSP-OLS images and adequate image analysis software such as ImageJ (in some cases to be complemented by further statistics software) provide a useful perspective for the analysis of spatial change. Since there is a stable and significant correlation between social and economic variables (population density, GDP PPP) and luminosity, such image analyses contain important information on spatial economic development. For micro-spatial comparisons, especially of urban areas, more precise results can be expected from using radiance-calibrated images. Analysis of night imagery is certainly not adequate to replace the statistical analysis of regional data, but it is a good tool to confirm and illustrate patterns of spatial heterogeneity and spatial dependence over time.

Further to that, the approach requires modest technical effort at no cost.

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