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Working Paper

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CESifo Working Paper, No. 6550

Provided in Cooperation with:

Ifo Institute – Leibniz Institute for Economic Research at the University of Munich

Suggested Citation: Bonilla, Roberto; Kiraly, Francis Z.; Wildman, John (2017) : Marriage Premium and Class, CESifo Working Paper, No. 6550, Center for Economic Studies and ifo Institute (CESifo), Munich

This Version is available at:

<https://hdl.handle.net/10419/167536>

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Impressum:

CESifo Working Papers

ISSN 2364-1428 (electronic version)

Publisher and distributor: Munich Society for the Promotion of Economic Research - CESifo GmbH

The international platform of Ludwigs-Maximilians University's Center for Economic Studies and the ifo Institute

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Marriage Premium and Class

Abstract

We present an equilibrium model with inter-linked labour and marriage markets, and argue that search frictions play an important role in explaining the male marriage premium. If men are viewed as bread-winners within households, this expectation affects their job search behaviour, which in turn can lead to marital wage gaps ranked across productivities. Male heterogeneity is neither necessary nor sufficient for this premium. Nonetheless, when coupled with female heterogeneity in the marriage market, male reservation wages linked to productivities do affect the resulting assortative matching structure, and equilibrium marriage classes exhibit their own patterns of marital wage differentials. Using UK data on wages, male/female heterogeneity and marital classes, we also carry out the first empirical test of the role of search frictions in generating the male marriage premium.

JEL-Codes: D830, J120, J640.

Keywords: frictional labour markets, frictional marriage markets, marriage premium, classes.

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21 June 2017

We are grateful for comments received from Jim Albrecht, Carlos Carillo-Tudela, Melvyn Coles, Marco Francesconi, Guido Menzio, Fabien Postel-Vinay, Alberto Trejos, Mich Tvede and Susan Vroman, as well as participants at the Search and Matching Annual Meeting (Aix-en-Provence, 2015), Econometric Society North American Summer Meeting (St. Louis, 2017), and seminar participants at SUNY Albany, Fribourg, Geneva, Munich and Neuchatel. Bonilla would also like to thank SUNY Albany and University of Oviedo for their hospitality while working on this project.

1 Introduction

Why do married men generally earn more than unmarried men? The question has intrigued the economics profession for a long time, as the so-called male marriage premium has consistently been found to be anywhere between 10% and 40% or even as high as 50%. In sharp contrast, the female marital earnings gap tends to be negative and much smaller in magnitude.

The existing literature offers two broad explanations which differ in terms of the direction of causality they favour. According to the *specialisation* theory developed by Becker (1973, 1974, 1985, 1991) and Mincer and Polachek (1974), marriage allows for economies of scale and intra-household specialisation of labour, with men being able to focus on formal employment, which in turn tends to enhance productivity and earnings. In a similar vein, Grossbard-Shechtman and Neuman (2003) note that the spouse also increases a married man's productivity directly by augmenting his human capital and general well-being.

In contrast, regarding to the so-called *selection* hypothesis, Nakosteen and Zimmer (1987), Davies and Peronaci (1997) and Gray (1997) argue that men who are more productive at work and/or earn high wages are more likely to get married. This may be the case if some traits that make men more productive in the labour market (such as loyalty, honesty and dependability) are also valued in the marriage market. Interestingly, this selection effect could also work indirectly through labour market gender discrimination, with employers favouring married men because they too associate marriage with productivity-enhancing traits.

This raises two important questions. First, why is there no female marriage premium? After all, there is no obvious reason why the logic of the above two explanations should not hold symmetrically for both sexes. Women are just as likely to have unobservable characteristics that are valued in employment as well as marriage, so why is the selection argument presumed one-sided? Similarly, unless there are explicit gender differences which stem from comparative advantages in human capital accumulation, labour market discrimination or intra-household bargaining positions, the specialisation argument is clearly also biased.

On the face of it, this unequal treatment might appear a reasonable ad hoc assumption, especially given the gap in observed male and female mar-

riage premia. In fact, however, one could argue that it obscures and overlooks a crucial insight: whatever the source of such asymmetry, the upshot is that men are seemingly regarded as providers in a household. Whether this is simply because of attitudes towards marriage, or it derives from family members' differing options in the labour market, it is this belief that ultimately guides the behaviour of marriage partners, under both approaches. We argue that such within-household expectations should be modelled explicitly, as they affect the entire job finding process, and subsequently the wages earned by men.

Second, it is apparent that current explanations of male marriage wage differentials follow the standard neo-classical approach which tends to equate human capital with earnings. But if one takes the job finding process seriously and considers search in frictional labour markets, productivity and wages can no longer be assumed fully inter-changeable. In turn, this means that the link between male productivity and marriage - as assumed by the specialisation and selectivity theories - may not be quite so straightforward either.

What is needed therefore is an equilibrium analysis, with male wage differentials as possible outcomes in a framework where labour market decisions are influenced by expectations and behaviour in the marriage market, and vice-versa. Naturally, this framework should also include search frictions, as it takes time, effort and luck to find a job and a suitable spouse.

We therefore construct a model where the two frictional markets are inter-linked. In the marriage market, men and women engage in two-sided sequential search for partners. Women prefer high earners (breadwinners), and this is common knowledge. In our setup, these marriage market expectations are captured quite neatly by women's reservation match value.

In turn, unemployed men engage in job search, knowing that their earnings determine whether or not they can get married. Formalising this sequential decision problem is, however, far from trivial, as one needs to capture the possibility of a man encountering wages that make him marriageable or not. To that end, we introduce and analyse what we call *constrained search* problem facing an unemployed male, whose access to the marriage market depends on his earned wage. This variant of the standard sequential job search constitutes the bridge between the two markets in the sense that the

optimal strategy (reservation wage) is now a function of women's reservation match value.

In a preliminary study, Bonilla and Kiraly (2013) show that a male marriage premium can indeed be an equilibrium outcome in a model where men are homogeneous in terms of productivity. Here, we expand that simple framework in two important ways. In order to gain a better understanding of the role of search frictions and productivity differences in generating patterns of marital wage premia, we first allow for male heterogeneity in the labour market. Following that, we also include female heterogeneity, with women being characterised by different levels of attractiveness in the marriage market. The setup with two-sided heterogeneity then provides a natural framework to examine the question of who marries whom.

We make several contributions. First, with one-sided heterogeneity, we confirm that a positive male marriage premium can occur in equilibrium. The key insight is that with job search, the optimal stopping strategy determines both the probability of getting married and the average wages earned by various groups of men.

We show that with constrained sequential job search, the optimal reservation wage may in fact be chosen so low that it precludes marriage. The logic is quite intuitive: although marriage does lead to an increase in utility for a man, the opportunity cost of holding out for a wage acceptable to women may outweigh this utility gain if the required well-paid jobs are relatively difficult to encounter. As a result, there will be single men on relatively low wages and men who marry *because* they were lucky and found good jobs.

Second, we establish that ex-ante productivity heterogeneity is neither a necessary nor a sufficient condition for the existence of male marriage premium. Such wage differentials can exist *within* each male productivity group, while the pattern of rankable marriage premia *across* productivity types does not exclude the possibility of zero marital earnings gaps for men with different productivities.

Third, we examine the endogenous formation of marital classes where the sorting occurs along male wages (and, to some extent, productivities) and female levels of attractiveness. Crucially, the distribution of wages earned by males is itself endogenous, determined by the optimal reservation wage strategies of different men. This is in contrast with the setup used in the

seminal Burdett and Coles (1997) paper and virtually all subsequent work on assortative matching in the presence of search frictions, which assume an exogenous distribution of wages.¹ We construct and establish the existence of a search equilibrium with positive assortative mating, fully characterise it in terms of marriage wage patterns across classes, and show that the bottom class will always contain men with zero marriage premium.

Fourth, we carry out the first empirical investigation of the role of search frictions in explaining the male marital earnings gap.² Our findings strongly indicate that marriage premia follow the patterns across male productivity types suggested by our model. Furthermore, in our sample we find evidence of assortative matching and a bottom marital class, where men with very low socio-economic characteristics are in partnerships with similarly "below average" women and indeed exhibit a zero marriage premium.

The literature on male marriage premium, excellently surveyed by Grossbard-Shechtman and Neuman (2003), Daniel (1995) and Hersch (2003), is almost exclusively empirical in nature. In the main, the studies use either cross sectional data or models with panel data, as well as some econometric models based on marriage market analysis. The standard methodology is to allow for individual-specific fixed effects in order to eliminate unobserved heterogeneity. Korenman and Neumark (1991), Cornwell and Rupert (1995) and Schoeni (1995) are good examples of this approach.

Although much of the empirical work appears to confirm the existence of a link between productivity differences and male marriage wage premium, the evidence in support of either the selection or specialisation explanations is frustratingly mixed.

The results of Chun and Lee (2001) suggest that the selection effect is minimal, while Ginther and Zavodny (2001) find that only up to 10% of wage premium is the result of selection. In contrast, Loh (1996) and Hersch and Stratton (2000) argue that the effect of a potential productivity increase after marriage is also quite weak. Blackburn and Korenman (1994) consider

¹The only other paper we are aware of that looks at endogenous marriage partner traits is Chiappori et al. (2009), which includes an ex-ante choice of investment in education.

²The need for such an empirical appraisal of the role of frictions in explaining male marriage premium is stressed in Ponthieux and Meurs (2014). Incidentally, Bonilla et al. (2016) also provide evidence of the link between market frictions and marriage premium, but their focus is on testing the search theoretical approach to the beauty premium.

evidence for both theories but conclude that neither seem to be sufficient explanations for the existence of marriage premium. More recently, Antonovics and Town (2004) use an US sample that includes twins, and reach the conclusion that the observed male marital premium cannot be explained by the selection hypothesis. In contrast, Bardasi and Taylor (2008) use UK panel data and remark that although household specialisation is a relevant contributing factor, the earnings gap is more likely to be the result of a selection effect.

It is apparent therefore that existing studies have not been able to settle the question, and the controversy goes on. Perhaps a new approach is needed. In an illuminating paragraph which we quote in full, Grossbard-Shechtman and Neuman (2003) highlight the problems facing empirical research on marriage premium and stress the need for a fresh angle: "What such fixed-effects models tell us is what part of the marital differentials are due to selectivity into marriage and what part are productivity-enhancing effects of marriage. However, the selectivity effect could be a breadwinner effect: It could be that an unmeasured motivation to marry had led men to enhance their labor market productivity and earnings, knowing that this would increase their chances of getting married. So identifying a sequence running from productivity to marriage does not prove that it is not marriage that makes men more productive. If selectivity effects are breadwinner effects, it is the prospect of marriage that makes men more productive."

In this context, we hope that the market equilibrium approach we propose, which incorporates search frictions as well as the breadwinner hypothesis, will prove to be a fruitful contribution to the debate about the sources of male marriage premium.

The rest of the paper is structured as follows. First, we describe the model with male heterogeneity. Section 3 establishes the existence of a search equilibrium with marriage premium and examines the role of male productivity differences. Section 4 considers the augmented framework with two-sided heterogeneity. There, we construct a search equilibrium with two marital classes characterised by specific marriage premium patterns. Following that, an empirical section provides evidence in support of our theory. Concluding remarks appear in Section 6.

2 The model

The economy consists of a continuum of women and men, all risk neutral. Time is continuous and all agents discount the future at rate r . Only steady state equilibria are considered.

Men enter the economy unemployed and single. We first consider the scenario with two types of men who differ in terms of productivity. The subscript i (where $i = L, H$) denotes low- and high productivity type, respectively. Let u_i denote the steady state measure of single unemployed men of type i .

In the labour market, unemployed men use costless random sequential job search, and contact with a firm occurs at rate λ_0 , identical for the two types. A type i unemployed draws an offer from the exogenous wage distribution $F_i(w)$ with support $[\underline{w}_i, \bar{w}_i]$, and we assume that $\underline{w}_L < \underline{w}_H$, $\bar{w}_L < \bar{w}_H$ and $F_L(w) > F_H(w)$. A single man who is employed at wage w has flow payoff w . There is no on-the-job search and no job destruction, so a man's wage remains constant throughout his working life.

All single men look for potential female marital partners. Women can observe a man's employment status, his productivity and earned wage. In the marriage market, meetings occur according to a quadratic matching function, where the number of meetings is proportional to the product of the measure of searchers on each side of the market. This (see Mortensen, 1982) gives rise to a Poisson meeting technology with an arrival rate that is a linear function of the relevant measure of participants on the opposite side of the market. For men of type i , we denote this arrival rate by λ_m^i . A married man earning wage w enjoys flow payoff $w + y$, where $y > 0$ captures the utility of marriage. Assume that y is the same for both types of men, and there is no possibility of divorce.

Let n denote the measure of single women, and let x denote the exogenous flow payoff of a single woman. Women do not look for jobs³, but use costless random sequential search to locate single males. Let λ_w^i and λ_w^u be the parameters of the Poisson process and hence the rate at which a woman meets a type i employed man or an unemployed, respectively. Upon contact, a woman decides whether to accept or reject a marriage proposal. Let N_i

³If they did, x could of course be endogenous, just as it would under any other scenario where women are active in the labour market.

denote the measure of employed single men of type i who are acceptable to women. A married woman's flow payoff is equal to her partner's wage earnings. Upon marriage a woman gives up x , so we assume that $x < \bar{w}_L$ in order to ensure that a marital partnership with either type of man does indeed have a positive potential surplus.⁴

Marriages are for life, but couples and singles alike leave the economy at an exogenous rate δ . Every time an unemployed single man of type i accepts a job, gets married or leaves the economy, he is replaced by another type i unemployed single man. Hence, u_i can be treated as exogenous. Similarly, we assume that a new single woman enters the economy every time a single woman gets married or exits, so n can also be regarded as exogenous.

Denote by λ the parameter which measures the efficiency of the meeting process. Let N_i denote the measure of employed marriageable single men of type i . Then, from a single man's point of view, the average instantaneous rate at which meetings occur is $\lambda_m^i = \frac{\lambda(N_H+N_L)n}{(N_H+N_L)} = \lambda n$. Similarly, from the point of view of single women, the rate at which meetings with unemployed men occur is $\lambda_w^u = \frac{\lambda(u_H+u_L)n}{n}$, while the corresponding rate of meeting employed men of type i is $\lambda_w^i = \frac{\lambda(N_H+N_L)n}{n} \frac{N_i}{(N_H+N_L)} = \lambda N_i$. The steady state measure N_i is of course endogenous, and consequently so is λ_w^i , together with the steady state measures of un-marriageable men of type i (bachelors, denoted by B_i) and of married men of type i (denoted by M_i).

Before we proceed with the analysis, we define the marriage premium among type i men (MP_i) as the difference between the average wage of type i married men (\tilde{w}_i^M) and the average wage of type i single men (\tilde{w}_i^S). Throughout, superscripts S and M stand for "single" and "married", respectively. Finally, in the context of the paper, "marriageable" and "eligible" are used interchangeably.

⁴Blundell et al. (2015) show that female attachment to the labour market weakens considerably after marriage. This has an immediate impact on their earnings (due to non-participation), but also a lasting impact in the form lower future wages and part time work even after reattachment to the labour market. In turn, Gould and Paserman (2003), Oppenheimer (1988), Oppenheimer and Lew (1995), Loughran (2002) and Loughran and Zissimopoulos (2009) provide evidence that women build this into their expectations and behaviour in the marriage market.

3 Marriage Premium

Given sequential search and the fact that utilities are increasing in wages, we show that women and men alike use optimal strategies characterised by the reservation property. However, crucially, men's best response is in fact a reservation wage *function* $R_i(T_i)$, where T_i is the reservation match of women.

3.1 Women

In this section, we first establish that women refuse to marry unemployed men of type i if the reservation match of females is high enough. To see that, consider a married and employed man of type i . Without the possibility of either on-the-job search or divorce, and conditional on survival, standard considerations give the discounted expected lifetime utility:

$$V_i^M(w) = \frac{w + y}{r + \delta}.$$

Although the above utility is clearly independent of productivity type, in the interest of clarity we will continue to use these subscripts when referring to this employment value.

Now consider a married but unemployed man of type i . If such a man receives a wage offer w and behaves optimally in each period, the relevant Bellman equation is:

$$v_i^M(w) = \max \{V_i^M(w), J_i^M\}$$

In the above, J_i^M denotes a jobless married man's value from continued job search, and it is given by:

$$J_i^M = \frac{1}{r + \delta + \lambda_0} \left[y + \lambda_0 \int_{\underline{w}_i}^{\bar{w}_i} v_i^M(w) dF_i(w) \right].$$

Above, the right-hand side captures the expected discounted return. Recall that y is the instantaneous utility from a marital partnership, while the second term in the brackets describes the expected utility of any change in the value of this man's state, computed as the probability he receives a job offer (λ_0) times the expected continuation value associated with the wage offer w , given that the offer can of course be rejected.

Note that J_i^M is independent of any current wage offer. In turn, $V_i^M(w)$ is continuous and increasing in the accepted offer w , so the optimal strategy has the reservation property. Denote the relevant reservation wage by R_i^M , and manipulate the above equation to obtain the familiar version:

$$(r + \delta)J_i^M = y + \lambda_0 \int_{R_i^M}^{\bar{w}_i} [V_i^M - J_i^M] dF_i(w).$$

Given the above, the married man's reservation wage R_i^M solves the standard equation:

$$J_i^M = V_i^M(R_i^M) = \frac{R_i^M + y}{r + \delta},$$

and we get:

$$R_i^M = \frac{\lambda_0}{r + \delta} \int_{R_i^M}^{\bar{w}_i} [w - R_i^M] dF_i(w).$$

Further integration by parts leads to:

$$R_i^M = \frac{\lambda_0}{r + \delta} \int_{R_i^M}^{\bar{w}_i} [1 - F_i(w)] dw \equiv \underline{R}_i \quad (1)$$

Importantly, $\underline{R}_H > \underline{R}_L$, because the two types of men face different prospects in the labour market.

Furthermore, \underline{R}_i is also the reservation wage of type i unemployed men in a standard sequential job search problem where there is no link with the marriage market. As we will show later, \underline{R}_i is in fact the lowest reservation wage for each type.

In principle, women could of course marry unemployed men too. Let us therefore examine the situation of a woman who is married to a jobless type i man. Introducing the subscript U_i to denote marriage to such an unemployed, her value function $W_{U_i}^M$ is given by:

$$(r + \delta)W_{U_i}^M = \lambda_0 \int_{\underline{R}_i}^{\bar{w}_i} [W_i^M(w) - W_{U_i}^M] dF_i(w),$$

where $W_i^M(w) = \frac{w}{r + \delta}$ is the discounted lifetime utility of being married to a type i employed man who earns wage w . The interpretation is straightforward: the flow value of being married to a type i unemployed man is

essentially the potential benefit from any increase in her utility if the man finds and takes up a job offer w .

For T_i to be a female reservation match, it needs to satisfy $W_{U_i}^M = \frac{T_i}{r+\delta}$. Making use of $W_i^M(w)$ above leads to:

$$(r + \delta)W_{U_i}^M = \frac{\lambda_0}{r + \delta} \int_{\underline{R}_i}^{\bar{w}_i} [w - T_i] dF_i(w).$$

Since the right-hand side is equal to \underline{R}_i for $T_i = \underline{R}_i$, we have:

$$(r + \delta)W_{U_i}^M = T_i = \underline{R}_i$$

On the other hand, for $\underline{R}_i < T_i$ we instead have:

$$(r + \delta)W_{U_i}^M > \underline{R}_i$$

Therefore, women reject marriage to a type i unemployed man if $T_i > \underline{R}_i$. For the rest of Section 3, we work under the assumption that this is indeed the case.⁵

Next, we derive the women's reservation match (wage) T_i .

First, note that when a woman turns down an employed man, she returns to the same pool of potential partners. From her point of view, the expected value of continued search is therefore independent of the wage she rejected, and reflects the mix of high- and low productivity men.

Second, as there is no job destruction and a married woman's instant utility is not affected directly by her partner's productivity, the lifetime utility of accepting a man with a wage w is simply the discounted value of that wage: $\frac{w}{r+\delta}$, which is clearly continuous and increasing in w .

Overall, this means that with sequential search, the optimal stopping strategy has the familiar reservation property: a woman will accept a type i man as long as his wage is at least as high as T_i .

From the point of view of a single woman, the key considerations are: the number of single marriageable employed men of each type (N_i), and

⁵By doing so, we eliminate the uninteresting (for the purpose of this section) equilibrium where the marriage market does not affect job search. However, we do allow for this in Section 4.

the wages earned by these eligible single men, captured by the distribution function $G_i(\cdot)$. Therefore, in order to compute a single woman's utility, it is helpful to determine (N_i) and $G_i(\cdot)$ first.

Eligible men of type i get married at rate λn and die at rate δ , while unemployed men of type i find marriageable wages at rate $\lambda_0 [1 - F_i(T_i)]$. Therefore, the steady-state equation for N_i is:

$$N_i(\lambda n + \delta) = u_i \lambda_0 [1 - F_i(T_i)],$$

which gives

$$N_i = \frac{u_i \lambda_0 [1 - F_i(T_i)]}{\lambda n + \delta} \quad (2)$$

In turn, $G_i(\cdot)$ is obtained from N_i above, together with the inflow-outflow equation:

$$u_i \lambda_0 [F_i(w) - F_i(T_i)] = (\delta + \lambda n) N_i G_i(w)$$

From here, we get:

$$G_i(w) = \frac{F_i(w) - F_i(T_i)}{1 - F_i(T_i)}. \quad (3)$$

Before we turn to the female reservation match, let us complete the picture by finding the steady state measures of bachelors (B_i) and married men (M_i), together with their respective wage distributions. Bachelors are men who have accepted jobs that make them un-marriageable, so their inflow-outflow equation is:

$$u_i \lambda_0 [F(T_i) - F(R_i)] = \delta B_i \quad (4)$$

Let $E_i(w)$ denote the distribution of wages amongst these men. We have:

$$u_i \lambda_0 [F(w) - F(R_i)] = \delta B_i E_i(w),$$

and therefore:

$$E_i(w) = \frac{F(w) - F(R_i)}{F(T_i) - F(R_i)}. \quad (5)$$

In turn, the steady state condition for the measure of married men is simply:

$$\lambda n N_i = \delta M_i \quad (6)$$

Finally, as the rate at which marriageable men get married and the rate at which married men exit the economy are both independent of their wages, the distribution of wages earned by type i married men is also $G_i(w)$.

We now turn to the reservation match set by women. Recall that the rate at which a woman meets an eligible single employed man of type i is $\lambda_i^w = \lambda N_i$. Denote the expected value of being single for a woman by $W^S(w)$. Then, provided $T_i > \underline{R}_i$ for both male types, we have:

$$\begin{aligned} (r + \delta)W^S(w) &= x + \lambda N_H \int_{T_H}^{\bar{w}_H} \left[\frac{w}{r + \delta} - W^S \right] dG_H(w) + \\ &\quad + \lambda N_L \int_{T_L}^{\bar{w}_L} \left[\frac{w}{r + \delta} - W^S \right] dG_L(w). \end{aligned}$$

Making use of (2) and (3), one obtains the alternative expression:

$$\begin{aligned} (r + \delta)W^S(w) &= x + \frac{\lambda u_H \lambda_0}{\lambda n + \delta} \int_{T_H}^{\bar{w}_H} \left[\frac{w}{r + \delta} - W^S \right] dF_H(w) + \\ &\quad + \frac{\lambda u_L \lambda_0}{\lambda n + \delta} \int_{T_L}^{\bar{w}_L} \left[\frac{w}{r + \delta} - W^S \right] dF_L(w) \end{aligned} \quad (7)$$

The left-hand side is the flow value for a single woman. On the right-hand side, we have the instantaneous utility x plus the expected increase in utility that results from any change in the value of her state, which is computed as the probability she meets an eligible type i employed man times the expected increase in value. Since $W^S(w) = T_i/(r + \delta)$, we have $T_H = T_L$, so from now on we denote the women reservation wage (match) as simply T .

Please observe that W^S , and therefore T itself are both independent of men's search strategy R_i . This is because the two distributions of posted wages are exogenous: type i men's reservation wage strategy does not affect $F_i(\cdot)$, and nor does it affect the rate at which job searchers find wages that ensure the prospect of marriage. Finally, as one can see from (3), the measure of eligible men (N_i) is also independent of R_i .

3.2 Men

Unemployed single men face a decision problem that we term *constrained job search*. These men sample wage offers knowing that they can only get married if they land a wage at least as high as women's reservation match. With sequential search, a type i unemployed single man uses an optimal stopping strategy that essentially boils down to having a reservation wage function $R_i(T)$. The shape of this best reply function is crucial for our analysis, and we show it has pretty robust characteristics.

As mentioned before, the more interesting and empirically relevant range is $T \in (\underline{R}_i, \bar{w}_i]$, where unemployed men are turned down by women. In contrast, employed single men could in principle be on wages that ensure marriageability.

Consider an unemployed single man of type i who has just obtained an offer w . This wage makes him either marriageable or not marriageable. Denote by $V_i^S(w)$ the value of accepting such a job. Then, standard considerations mean that:

$$V_i^S(w) = \begin{cases} \frac{w}{r + \delta} + \frac{\lambda n}{(r + \delta + \lambda n)(r + \delta)} y & \text{if } w \geq T \\ \frac{w}{r + \delta} & \text{if } w < T \end{cases} \quad (8)$$

Clearly, the value of accepting an offer w is a piecewise linear function: it is increasing in wage w and is discontinuous at $w = T$.

Let $\bar{v}_i(w)$ be the expected value for a type i unemployed man who currently has an offer w that confers marriageability, and let $\underline{v}_i(w)$ denote the value of currently holding a wage offer that precludes marriage. This man has to decide whether to accept or reject such a job offer. Let J_i denote the value that a type i unemployed man attaches to the option of continued job search, and note that for any currently held wage offer w , this value J_i is in fact independent of w .

Then, the corresponding *two* Bellman equations are:

$$\bar{v}_i(w) = \max \{ V_i^S(w \mid w \geq T), J_i \}$$

and

$$\underline{v}_i(w) = \max \{ V_i^S(w \mid w < T), J_i \}$$

Note that for $y > 0$, we have $V_i^S(w \mid w \geq T) > V_i^S(w \mid w < T)$, and therefore $\bar{v}_i \geq \underline{v}_i$. The latter is with strict inequality if $V_i^S(w \mid w \geq T) > J_i$, which in turn is true for $T > \underline{R}_i$ (see the proof of Proposition 1 below).

Independently of a current wage offer w , and of course provided there is contact with another firm, continued job search leads to a new wage offer w' . Given $F_i(\cdot)$, the probability with which any such w' is higher (or lower) than the female reservation wage (match) T depends only on T itself.

Importantly, all the above means that J_i is a *convex combination* of $\bar{v}_i(w)$ and $\underline{v}_i(w)$:

$$J_i = \frac{\lambda_0}{r + \delta + \lambda_0} \left[\int_T^{\bar{w}_i} \bar{v}_i(w') dF_i(w') + \int_{\underline{w}_i}^T \underline{v}_i(w') dF_i(w') \right]. \quad (9)$$

A given wage offer either confers marriageability or precludes the prospect of marriage. The right-hand side captures the discounted expected return to search: the terms in the big brackets describe the expected utility of any change in the value of a type i unemployed man's state. Here, this is computed as the probability that he receives a wage offer (λ_0), times the expected continuation value associated with the offer w' , bearing in mind not only that an offer can of course be accepted or rejected, but also that if accepted, this wage then either guarantees or jeopardises completely one's marriage prospects.

The above continuation value J_i is continuous for any $T < \bar{w}_i$, and is decreasing in T for $\bar{v}_i > \underline{v}_i$. In contrast, J_i is independent of T for $T \geq \bar{w}_i$, since for such T 's $F_i(T) = 1$. With this in mind, let \hat{T}_i be the T for which $J_i(T) = \frac{w}{r+\delta}$ at $w = T$. This particular female reservation match will prove extremely useful throughout our analysis below.

Making use of all the above, we are ready to fully characterise the male reservation wage function for any T .

Proposition 1 (A) *For $T \in (\underline{R}_i, \bar{w}_i]$, the optimal strategy has the reservation property. The reservation wage function $R_i(T)$ solves*

$$R_i(T) = \min \{w : V_i^S(w) \geq J_i\},$$

it is continuous and piecewise differentiable, with:

(i) $R_i = T$ for $T \in (\underline{R}_i, \hat{T}_i]$;

(ii) $R_i < T$ and decreasing in T for $T \in (\hat{T}_i, \bar{w}_i]$;

(B) For $T \leq \underline{R}_i$ and $T \geq \bar{w}_i$ the optimal strategy has the reservation property, and the reservation wage is \underline{R}_i .

Proof. *Part (A).* To address (i), recall that \hat{T}_i is the T for which $J_i(T) = \frac{w}{r+\delta}$ at $w = T$. If $T = \hat{T}_i$, the optimal policy is a reservation wage R_i and $R_i = T$ as $V_i^S(w \mid w < T) < J_i$ for $w < \hat{T}_i$ and $V_i^S(w \mid w \geq T) > J_i$ for $w \geq \hat{T}_i$. As T decreases, $J_i(T)$ increases and $\frac{T}{r+\delta}$ decreases. For a small decrease in T , we have $J_i > V_i^S(w \mid w < T)$ and $J_i < V_i^S(w \mid w \geq T)$, so the reservation wage equals T . Let \tilde{T}_i denote the T so low that $J_i(\tilde{T}_i) = V_i^S(w \mid w = \tilde{T}_i)$. It is easy to show that $\tilde{T}_i < \underline{R}_i$, and therefore $V_i^S(w \mid w \geq T) > J_i$ for $T \in [\underline{R}_i, \hat{T}_i]$. It follows that $R_i = T$ for $T \in (\underline{R}_i, \hat{T}_i]$.

To address (ii), start at $T = \hat{T}_i$, where $\frac{T}{r+\delta} = J_i(T)$, and consider an increase in T . Then, $\frac{T}{r+\delta}$ increases and $J_i(T)$ decreases. Since the discontinuity in $V_i^S(w)$ moves rightward, it follows that $J_i = V_i^S(w \mid w < T)$ holds for a wage lower than T . This is our reservation wage R_i , which clearly decreases as T increases. The derivation of both R_i and \hat{T}_i are relegated to Appendix A, where we also show that given a finite y and finite job offer arrival rate, the $R_i < T$ range always exists. One can also conclude that the optimal reservation wage is continuous in T , with a corner at \hat{T}_i . Finally, for $T = \bar{w}_i$, we have $F_i(T) = 1$ and the problem reverts to one of standard sequential job search.

Part (B). From (1), we know that once married, a type i unemployed man drops his reservation wage to \underline{R}_i . Despite that, if $T \leq \underline{R}_i$ a woman would still be willing to marry such a man. Recall that $J_i^M = \frac{\underline{R}_i + y}{r+\delta}$ and let J'_i denote the value of a single unemployed man in this scenario. For any optimally chosen reservation wage R'_i that is higher than (or equal to) T we have $V_i^S(w \mid w \geq T)$ for any acceptable offer, and $V_i^S(w \mid w \geq T)$ continuous in w . One can therefore write J'_i using the standard asset pricing equation:

$$(r + \delta) J'_i = \lambda n [J_i^M - J'_i] + \lambda_0 \int_{R'_i}^{\bar{w}_i} \left[\frac{w}{r + \delta} + \frac{\lambda n}{(r + \delta)(r + \delta + \lambda n)} y - J'_i \right] dF_i(w),$$

with the reservation wage R'_i solving $V_i^S(R'_i) = J'_i$. It is easy to show that \underline{R}_i is indeed higher than T and satisfies $V_i^S(\underline{R}_i) = J'_i$, so $R'_i = \underline{R}_i$.

Finally, for $T \geq \bar{w}_i$, we have $F_i(T) = 1$ and the problem reverts to one of standard sequential job search. ■

The reservation wage functions of unemployed men are illustrated in Figure 1. The diagram shows that men's best response functions are non-monotonic and attain their respective maximum values at $T = \hat{T}_i$, with $\hat{R}_i = R_i(\hat{T}_i) = \hat{T}_i$. Since $F_H(w) < F_L(w)$, we have that $\underline{R}_L < \underline{R}_H$ and $\hat{T}_H > \hat{T}_L$. Clearly, $\hat{R}_i > \underline{R}_i$.

This non-monotonicity captures an interesting trade-off faced by all men, for different levels of T . When the marriage problem is non-trivial, an increase in the male reservation wage has two effects. On the one hand, it guarantees marriageability - if it satisfies women's reservation match. On the other hand, any increase in reservation wage comes at the cost of narrowing ones job prospects.

If the female reservation match is relatively low (but still above \underline{R}_i), the opportunity cost is not too high, and therefore men hold out for such a wage. However, this opportunity cost increases with T and there is a threshold value (\hat{T}_i) for which the negative effect is no longer compensated by the prospect of marriage.

For $T \in (\hat{T}_i, \bar{w}_i)$, we can show (see Appendix A) that the reservation wage solves:

$$R_i = \frac{\lambda_0}{r + \delta} \left[\int_{R_i}^{\bar{w}_i} [1 - F_i(w)] dw + \frac{[1 - F_i(T)] \lambda n}{r + \delta + \lambda n} y \right] \quad (< T)$$

By setting a reservation wage lower than what women expect, a single man risks throwing away the prospect of marriage. Nonetheless, he is willing to accept this gamble purely because of search frictions and what one might call the "bird in hand" effect. Now, a job offer is deemed acceptable by an unemployed single man even if it precludes marriage, given the relatively high threshold set by women. Nonetheless, such a wage may still be sufficiently high so that an unemployed turns down the option of holding out for even higher (but unlikely) offers.

For higher and higher female reservation matches the likelihood of encountering such increasingly high wages decreases further, since $1 - F_i(T)$ decreases, and with it, men's reservation wage. Reservation wages gradually fall, to the point where men give up completely and rely on pure luck - after all, even the man who sets the lowest possible reservation wage may still be able to get married if he accidentally lands a good job.

3.3 Equilibrium with Marriage Premia

A search equilibrium is a system $\{R_i, T, N_i, G_i(\cdot), B_i, E_i(\cdot), M_i\}$, where the reservation wage function is summarised in Proposition 1, T is obtained from (7), while the various steady state measures of men, together with the distributions of earned wages result from inflow-outflow equations (2) - (6).

Recall that the marriage premium for type i men (MP_i) has been defined as the difference between the average wage earned by type i married men and the average wage of type i single men. Denoting the former by \tilde{w}_i^M and the latter by \tilde{w}_i^S , we simply have $MP_i \equiv \tilde{w}_i^M - \tilde{w}_i^S$.

An equilibrium with zero marriage premium for all men is less relevant from an empirical point of view. Therefore, we focus on equilibria that exhibit a positive marital wage premium for at least one type. Crucially, such equilibria allow us to compare marriage premia *across* types.

Note that in the context of our model, to obtain a measurable marriage wage differential in the first place it must be true that at least some men can indeed get married. We work under the assumption that this is the case, so $T < \bar{w}_i$ for all i . Also, let $\hat{T}_H < \bar{w}_L$ for now. Straight away, one can establish the link between reservation wages and marital wage differences.

Lemma 1 $MP_i > 0$ for $R_i < T$, and $MP_i = 0$ for $R_i = T$.

Proof. Note that

$$\tilde{w}_i^S = \frac{B_i}{B_i + N_i} \int_{R_i}^T w dE_i(w) + \frac{N_i}{B_i + N_i} \int_T^{\bar{w}_i} w dG_i(w)$$

and

$$\tilde{w}_i^M = \int_T^{\bar{w}_i} w dG_i(w).$$

For $T = R_i$, we have $B_i = 0$ and $E_i(w)$ collapses to zero, so $\tilde{w}_i^M = \tilde{w}_i^S$. On the other hand, for $T > R_i$, we have $B_i > 0$, and hence $\tilde{w}_i^M > \tilde{w}_i^S$. ■

Please note the importance of the above result: a positive marriage premium is possible only if the region where $R_i < T$ exists. In turn, as shown above, this can occur only when there are search frictions in the labour market. Obviously, marriage market search frictions are also needed if one is to talk about a meaningful reservation match T .

Theorem 1 Let \hat{x}_i such that $T(\hat{x}_i) = \hat{T}_i$, \underline{x} such that $T(\underline{x}) = \underline{R}_H$, and \bar{x} such that $T(\bar{x}) = \bar{w}_L$.

Then, for $x \in [\underline{x}, \bar{x}]$, there always is a non-empty set of parameter values for which a unique equilibrium exists where only employed men can get married and both productivity types do. The potential equilibria are as follows:

- (a) An equilibrium with $MP_L > MP_H > 0$ obtains iff $x \in (\hat{x}_H, \bar{x})$;
- (b) An equilibrium with $MP_L > MP_H = 0$ obtains iff $x \in (\hat{x}_L, \hat{x}_H)$;
- (c) An equilibrium with $MP_i = 0$ obtains iff $x \in (\underline{x}, \hat{x}_L)$.

Proof. We have established that R_H and R_L are continuous and non-monotonic, and it is easy to show that $\underline{R}_i < \hat{T}_i < \bar{w}_i$. Furthermore, we know that T is independent of R_i . Together, these guarantee the existence of an unique equilibrium.

Since $\partial T / \partial x > 0$ and $\underline{R}_H < \hat{T}_L < \hat{T}_H$, we have $\underline{x} < \hat{x}_L < \hat{x}_H < \bar{x}$. For $x \in [\underline{x}, \bar{x}]$, we have $\underline{R}_i < T < \bar{w}_L$ and therefore in an equilibrium only employed men can get married, and both L and H types do so as long as they earn a high enough wage.

For $\hat{x}_H < x < \bar{x}$ we have $\hat{T}_H < T < \bar{w}_L$, $R_H > R_L$ and $T > R_i$ for all i . From Lemma 1, $T > R_i$ implies there is a marriage premium for both types. Furthermore, $R_H > R_L$ means that, for a given T , the marriage premium is lower for H type men.

For $\hat{x}_L < x < \hat{x}_H$ we have $\hat{T}_L < T < \hat{T}_H$, $R_H \geq R_L$, $T > R_L$ and $T = R_H$. By Lemma 1, $T > R_L$ implies that there is a marriage premium for L type men but, as $T = R_H$, the H types do not exhibit a marriage premium.

Finally, for $\underline{x} < x < \hat{x}_L$ we have $\underline{w}_H < T < \hat{T}_L$ and $R_i = T$. Using Lemma 1 again, this means no marriage wage premium for either type. ■

Figure 1 plots T and R_i and captures equilibrium (a). The intersection point between the relevant best response functions captures the equilibrium where both productivity types exhibit (positive) male marriage wage premium.

FIGURE 1 ABOUT HERE

Theorem 1 contains very strong results which predict that the marriage premium among more productive men is lower than that among less productive men. Such an equilibrium outcome is possible (and hence observable) only if there are search frictions in the labour market, which are necessary for the non-monotonicity of the reservation wage function.

For completeness, note that if $\hat{T}_H > \bar{w}_L$, this only restricts the set of equilibria in which both types of men can marry in principle, but otherwise our qualitative results carry through. The full round-up of the potential equilibria in this case is relegated to Appendix B.

4 Class

In what follows, we extend the model in three directions. First, we introduce a continuum of male types. Second, we consider female heterogeneity in the marriage market. Third, we also allow for the scenario where women are willing to marry unemployed men. These additional features allow for a richer set of results: male marriage premium patterns are now equilibrium outcomes that result not only from optimal constrained sequential job search, but also from a process of endogenous marital class formation.

To that end, assume that women are characterised by a measure of attractiveness y in the marriage market, with all men agreeing on the ranking. One can think of this as any non wage-related quality that is valued by men. Let y be distributed according to the cumulative distribution function $H(y)$ with support $[y, \bar{y}]$.

Similarly, men are indexed by type p , and each type p has a mass u .⁶ Let p take values between \underline{p} and \bar{p} . Men of type p face an exogenous wage distribution $F_p(\cdot)$ with support $[\underline{w}(p), \bar{w}(p)]$, with $\underline{w}(p) = p$ and $\bar{w}(p) = p + z$, where $z > 0$. Another natural assumption is that $F_p(w) < F_{p'}(w)$ for $p > p'$. Given this, one can associate higher type p with higher productivity.

Men and women choose reservation matches and subsequently marriage classes form endogenously. To keep the analysis simple, we construct an equilibrium with two classes only. That is, the population of men and women will endogenously divide into two sets: men from male Set 1 marry women from the female Set 1, and men from male Set 2 marry women from the female Set 2. Following that, Class 1 will contain highly desirable women paired up with very productive and/or high earning men, while Class 2 will have couples made up of less desirable women matched with less productive

⁶Having equal measures of each type amounts to having a uniform distribution of types, with density $\frac{1}{\bar{p} - \underline{p}}$.

and/or low earning men. The reservation match strategies of men and women are derived in detail below, but for now, a quick overview is probably helpful.

Following our earlier analysis, men who ignore the marriage market (either because they can, or because they have no choice) set a reservation wage $\underline{R}(p)$.

Employed or unemployed men from Set 1 who are accepted by all women choose reservation match y_1 . In turn, employed or unemployed men from Set 2 who are only accepted by women from female Set 2 choose reservation match y_2 . These reservation values are such that $y_1 \in [\underline{y}, \bar{y}]$ and $y_2 < \underline{y}$. As we do not rule out unemployed men being able to get married, let $y^*(p)$ denote the reservation match of a type p unemployed.

Similarly, the measure n_1 of Set 1 women choose reservation wage T_1 , while women from Set 2 (measure n_2) choose reservation wage T_2 . We limit our analysis to the case where $T_1 < \bar{p} + z$, and note at this point that n_1 and n_2 are of course endogenously determined.

To keep the dynamics simple, we assume that a y type woman is replaced by another woman of the same type whenever she gets married or exits the market. As a consequence, we can treat the total measure and distribution of single women as exogenous and are able to normalise it to 1 without any loss of generality. Hence, the measure of women in Set 1 (n_1) is simply $[1 - H(y_1)]$, while the measure of Set 2 women (n_2) is $H(y_1)$.

Given the reservation wage strategies $R(p)$, as well as $\underline{R}(p)$, T_1 and T_2 as defined above, the male population will split into groups, based on type and wage. More specifically, define threshold productivities p_1, p_2, p_3 such that:

$$\underline{R}(p_1) = T_1 \tag{10}$$

$$\hat{T}(p_2) = T_1 \tag{11}$$

and

$$p_3 = T_1 - z \tag{12}$$

Then, we obtain the following groups of men:

(i) Men with $p \in [p_1, \bar{p}]$ can afford not to care about the marriage market. These men can marry anyone, even when unemployed, and for them it is also true that $y^* = y_1$.

(ii) Men with $p \in [p_2, p_1)$ set $R(p) = T_1$ because $\hat{T}(p_2) = T_1$. These men can - once they get a job - marry women from Set 1. While unemployed, they may be willing to marry a woman from Set 2 if her y is high enough: that is, higher than $y^*(p)$.

(iii) Men with $p \in [p_3, p_2)$ will chance it in the sense that they set $R(p) < T_1$. If they are lucky and land a wage $w \geq T_1$, they subsequently have access to women from Set 1.

(iv) Men below p_3 can not hope to marry women from Set 1, since a p_3 type man has $\bar{w}(p_3) = T_1$. Their only chance is women from Set 2. Note that for these men to exist, we require that $\underline{p} < p_3$, so we work under this scenario.

Next, define p^* such that a Set 2 woman is indifferent between staying single or marrying an unemployed p^* man. With $p^* < p_3$, Set 2 women are willing to marry at least some men with $p < p_3$. Indeed, if $p^* < \underline{p}$, Set 2 women will in fact marry *all* unemployed men with $p < p_3$. As we will see, this condition ensures there is a second class in which some male types (those with $p < p_3$) have zero marriage wage premium.

Please note that constructing such a bottom class is by no means trivial. Firstly, this is because both p^* and p_3 are in fact endogenous. Secondly, women from Set 2 could in principle find partners from *three* groups of single men: those with $p \in [\underline{p}, p_3)$, those with $p \in [p_3, p_2)$ who are chancing it and are either unemployed or unlucky, as well as those who, while holding out for a wage that ensures marriageability to a Set 1 woman, may nonetheless be willing to marry some Set 2 women instead. Indeed, a Set 2 woman with attractiveness y will be accepted by men with productivity no higher than p' , where $p' \in [p_2, p_1)$ is defined as the productivity p for which $y^*(p) = y$.

4.1 Search

In this section we characterise the optimal reservation wage and/or match strategy for men, together with the optimal reservation match strategy for women. We assume $\bar{p} < T_1 < \bar{p} + z$.

4.1.1 Men

Employed men with $w \geq T_1$ These are single men with $p \in [p_3, \bar{p}]$, and the pool contains those who are very high type (so can ignore the marriage market), those who held out for a wage no lower than T_1 , and those who chanced it but were lucky and obtained a marriageable wage. Having access to all women, these men only accept women with $y \geq y_1$.

Let $V_1^S(w)$ denote their value of being single and employed at wage $w \geq T_1$, and $V^M(w, y)$ denote their value of being married to a type y woman. Then, $V^M(w, y) = \frac{w+y}{r+\delta}$ and:

$$(r + \delta)V_1^S(w) = w + \lambda n_1 \int_{y_1}^{\bar{y}} [V^M(w, y) - V_1^S(w)] dH'(y), \quad (13)$$

where:

$$H'(y) = \frac{H(y) - H(y_1)}{1 - H(y_1)} \quad \text{and} \quad n_1 = 1 - H(y_1). \quad (14)$$

The interpretation is straightforward: a single man who earns a wage w meets a desirable woman from Set 1 at rate λn_1 and gets married. Alternatively, one can rewrite the above as:

$$V_1^S(w) = \frac{w}{r + \delta} + \frac{\lambda}{[r + \delta + \lambda(1 - H(y_1))](r + \delta)} \int_{y_1}^{\bar{y}} y dH(y) \quad (15)$$

From the definition of a reservation match, we have:

$$V^M(w, y_1) = V_1^S(w)$$

Using this, we obtain:

$$y_1 = \frac{\lambda}{r + \delta} \int_{y_1}^{\bar{y}} [1 - H(y)] dy \quad (16)$$

The interpretation is standard, as the right-hand side reflects the opportunity cost of continued search in the marriage market.

Employed men with $w < T_1$ These are single men with $p \in [\underline{p}, p_2)$, and the group includes men with $p \in [p_3, p_2)$ - those who chanced it but were unlucky and ended up with an unappealing wage, as well as those with $p < p_3$ - men who were hopeless to start with. All these men can only marry women with attractiveness y lower than y_1 . By construction, they will accept them all, so there is no *third* class. This means that $V^M(w, \underline{y}) \geq V_2^S(w)$, where $V_2^S(w)$ denotes the value of being single for these men, obtained from:

$$(r + \delta)V_2^S(w) = w + \lambda n_2 \int_{y_2}^{y_1} [V^M(w, y) - V_2^S(w)] dH''(y), \quad (17)$$

where:

$$H''(y) = \frac{H(y)}{H(y_1)} \quad \text{and} \quad n_2 = H(y_1) \quad (18)$$

Alternatively, we can write:

$$V_2^S(w) = \frac{w}{r + \delta} + \frac{\lambda}{[r + \delta + \lambda H(y_1)](r + \delta)} \int_{\underline{y}}^{y_1} y dH(y) \quad (19)$$

Again, equating $V_2^S(w)$ and $V^M(w, y_2)$, one obtains the reservation match implicitly:

$$y_2 = \frac{\lambda}{r + \delta} \int_{y_2}^{y_1} [H(y_1) - H(y)] dy \quad (20)$$

Given the above, in order to satisfy $y_2 < \underline{y}$, we need:

$$\underline{y} > \frac{\lambda}{r + \delta + \lambda H(y_1)} \int_{\underline{y}}^{y_1} y dH(y)$$

Unemployed men These are single men who choose a reservation match $y^*(p)$ in the marriage market as well as a reservation wage $R(p)$ in the labour market.

From our previous analysis of the constrained sequential job search problem, we know there are two important ranges of relevant reservation wages.

(a) *Unemployed men with $p \in [p_2, p_1)$ who choose $R(p) = T_1$.*

Recall that as soon as any unemployed gets married, he subsequently sets $R(p)$ equal to $\underline{R}(p)$. As a consequence, Set 1 women do not marry them in the first place. Nonetheless, while unemployed, these men do set a reservation match $y^*(p)$ which may or may not be lower than y_2 .

Denote the value of unemployment of these men by $U_1(p)$, and also note that, once they find a wage higher than or equal to T_1 , their value of unemployment is lower than the value of employment: $U_1(p) < \frac{T_1}{(r+\delta)}$. After finding such a wage, their reservation match is set at y_1 and therefore while unemployed, their reservation match $y^*(p)$ must be less than y_1 .

Since unemployed men drop their reservation wage to $\underline{R}(p)$ after marriage, their value of being jobless but married is:

$$V^M(\underline{R}(p), y) = \frac{\underline{R}(p) + y}{r + \delta}$$

The discounted lifetime utility of being married while unemployed reflects the fact that once a man gets married his reservation wage drops to $\underline{R}(p)$.

In turn, their value of being unemployed is:

$$\begin{aligned} (r + \delta)U_1(p) = & \lambda_0 [1 - F_p(T_1)] \int_{T_1}^{\bar{w}(p)} [V_1^S(w) - U_1(p)] dF_p(w|w > T_1) + \\ & + \lambda n_2 [H(y_1) - H(y^*(p))] \int_{y^*(p)}^{y_1} [V^M(\underline{R}(p), y) - U_1(p)] dH(y|y_1 > y > y^*(p)) \end{aligned}$$

As mentioned, these unemployed men are not acceptable to Set 1 women. Nonetheless, they can of course look for a job and only accept it if the wage offered is above T_1 - in which case they become marriageable to Set 1 women. Alternatively, they could meet a Set 2 woman who will accept them despite the fact that they are unemployed. However, these men will only accept such women if in turn they have a sufficiently high y . This particular reservation match $y^*(p)$ solves:

$$V^M(\underline{R}(p), y^*(p)) = U_1(p)$$

From the definition of a reservation wage, a jobless man is indifferent between continued job search while single and marriage to a y^*

Next, using $V_1^S(w) = V^M(w, y_1)$ together with $U_1(p) = V^M(\underline{R}(p), y^*(p))$, and integrating by parts, we obtain y^* implicitly:⁷

$$\begin{aligned} y^*(p) = & \frac{\lambda_0}{r + \delta} \int_{T_1}^{\bar{w}(p)} [1 - F_p(w)] dw + \\ & + \frac{\lambda_0}{r + \delta} [1 - F_p(T_1)] [y_1 - y^*(p)] + \\ & + \frac{\lambda n_2}{r + \delta} \int_{y^*(p)}^{y_1} [H(y_1) - H(y)] dy - \underline{R}(p) \end{aligned} \quad (21)$$

The male reservation match has to compensate for: (i) the standard option of continued search for better wages while unemployed; (ii) the possibility of an utility gain (in terms of y) from potentially finding a wage higher than T_1 that allows a partnership with a Set 1 woman; and (iii) the option of continued search for women with higher y . These three effects are mitigated by the option of dropping the reservation wage to $\underline{R}(p)$ if in the position to get married while unemployed.

We can further characterise this reservation match $y^*(p)$. First, a simple implicit derivation of the y^* equation shows that $\partial y^*(p)/\partial p > 0$. Second, one can show that if $y_2 \leq \underline{y}$, then $y^*(p) \leq \underline{y}$ for $p \in (p_3, p_2]$. To see this, consider men with $p \in (p_3, p_2)$. Denote their value of unemployment by $U_2(p)$, and observe that $V_2^S = \frac{R(p)}{r + \delta} = U_2(p)$. If y_2 is their reservation match while employed (that is, when enjoying utility V_2^S), then it is also y_2 while jobless (that is, when enjoying utility $U_2(p) = V_2^S$).

Note that in principle, $y^*(p)$ can be higher or less than \underline{y} . To explore this further, define p' such that $y^*(p') = \underline{y}$. From the above, we know that $p' > p_2$.⁸ For $p' > p_1$, all men with $p \in (p_2, p_1)$ set $y^*(p) < \underline{y}$ and are willing to

⁷We present the implicit version as it has a more intuitive interpretation, but note that $y^*(p)$ can easily be derived explicitly.

⁸For $p = p_2$ we have $R(p_2, T_1) = T_1$, so it is true that $V_1^S(R(p_2)) = \frac{R(p_2)}{r + \delta} = U(p_2)$. The reservation match of these men is $y_2 < \underline{y}$.

marry all women. For $p' < p_1$, men with $p \in (p_2, p')$ also set $y^*(p) < \underline{y}$, while men with $p \in (p', p_1)$ set $y^*(p) > \underline{y}$ and only accept women with $y > y^*(p')$, where $y^*(p') > \underline{y}$.

(b) *Unemployed men with $p \in [p_3, p_2)$ who choose $R(p) < T_1$.*

These men are willing to accept wages that put them in the male Set 2. Of course, they may still be lucky and find wages that get them into Set 1, giving access to Set 1 women. Their value of unemployment is $U_2(p)$, where $(r + \delta)U_2(p) =$

$$\begin{aligned} \lambda_0 [F_p(T_1) - F_p(R(p))] \int_{R(p)}^{T_1} [V_2^S(w) - U_2(p)] dF_p(w|R(p)) &< w < T_1) + \\ + \lambda_0 [1 - F_p(T_1)] \int_{T_1}^{\bar{w}(p)} [V_1^S(w) - U_2(p)] dF_p(w|w > T_1) &+ \\ + \lambda n_2 \int_{\underline{y}}^{y_1} [V^M(\underline{R}(p), y) - U_2(p)] dH(y|y < y_1) \end{aligned}$$

These men could: *a)* accept a wage less than T_1 that puts them in Set 2, *b)* find a job that puts them in Set 1, or *c)* find an acceptable Set 2 woman and get married. The latter implies that for these men, $y^*(p) \leq \underline{y}$: jobless or not, they find Set 2 women desirable. Having accepted $R(p)$, they set $y^*(p) \leq \underline{y}$, and they are indifferent between unemployment and employment at wage $R(p)$.

Once again, recall that if a jobless man gets married, subsequently his reservation wage drops to $\underline{R}(p)$. Hence, the value of being married while unemployed is again $V^M(\underline{R}(p), y) = \frac{\underline{R}(p) + y}{(r + \delta)}$. From the definition of a reservation wage, and using $V_1^S(w)$ and $V_2^S(w)$, as well as $n_2 = H(y_1)$, we obtain:

$$\begin{aligned}
R(p) &= \frac{\lambda_0}{r + \delta + \lambda H(y_1)} \int_{R(p)}^{\bar{w}(p)} [1 - F_p(w)] dw + \\
&+ \frac{\lambda_0(r + \delta)}{r + \delta + \lambda H(y_1)} [1 - F_p(T_1)] [V_1^S(w) - V_2^S(w)] + \\
&+ \frac{\lambda}{r + \delta + \lambda H(y_1)} \int_{\underline{y}}^{y_1} [\underline{R}(p) + y] dH(y) \quad (< T_1).
\end{aligned} \tag{22}$$

The terms on the right hand side reflect that the reservation wage has to compensate for: (i) the standard option of continued search for better wages, (ii) the foregoing of the potential to marry a Set 1 woman (and the acceptance of only being able to marry Set 2 women), as well as (iii) the option of marrying a Set 2 woman while unemployed.

(c) *Unemployed men with $p \in [p_1, \bar{p}]$ and unemployed men with $p \in [\underline{p}, p_3]$.*

Recall that p_1 is such that $\underline{R}(p_1) = T_1$. Therefore, men with $p > p_1$ need not worry about marriage, since $\underline{R}(p) > T_1$, so they set their reservation wage equal to $\underline{R}(p)$.

4.1.2 Women

Women in Set 1 These women only marry men whose wages are higher than T_1 . The pool of such eligible singles is made up of the following:

- i) a measure N_p of single employed men with $p > p_1$, whose wages are higher than T_1 ;
- ii) a measure $u[\bar{p} - p_1]$ of single unemployed men with $p > p_1$, who have reservation wages $\underline{R}(p) > T_1$;
- iii) a measure N'_p of single employed men with $p \in [p_2, p_1)$, who have held out for a wage at least as high as T_1 ;
- iv) a measure N'_p of single employed men with $p \in [p_3, p_2)$, who chanced it, were lucky and are now earning a wage $w \geq T_1$.⁹

⁹The last two measures are equal because the wage distributions are exogenous.

Using steady-state equations analogous to the ones in Section 3, we obtain the measures and distribution of earned wages for eligible men ¹⁰:

$$G_p(w) = \frac{F_p(w) - F_p(T_1)}{1 - F_p(T_1)} \quad (23)$$

$$N_p = \frac{u\lambda_0[1 - F_p(\underline{R}(p))]}{\lambda n_1 + \delta} \quad (24)$$

$$N'_p = \frac{u\lambda_0[1 - F_p(T_1)]}{\lambda n_1 + \delta} \quad (25)$$

Given the above, for a woman in Set 1 the value of being single is:

$$\begin{aligned} (r + \delta)W_1^S &= \Psi \int_{p_1}^{\bar{p}} N_p \left[\int_{\underline{R}(p)}^{\bar{w}_p} \left(\frac{w}{r + \delta} - W_1^S \right) dG_p(w) \right] dp + \\ &+ \Psi u [\bar{p} - p_1] \int_{p_1}^{\bar{p}} \left[\frac{\underline{R}(p)}{r + \delta} - W_1^S \right] dp + \\ &+ \int_{p_2}^{\min(p_1, \bar{p})} N'_p \left[\int_{T_1}^{\bar{w}_p} \left(\frac{w}{r + \delta} - W_1^S \right) dG_p(w) \right] dp + \\ &+ \int_{p_3}^{\min(p_2, \bar{p})} N'_p \left[\int_{T_1}^{\bar{w}_p} \left(\frac{w}{r + \delta} - W_1^S \right) dG_p(w) \right] dp, \end{aligned} \quad (26)$$

where $\Psi = 1$ for $\bar{p} > p_1$ and $\Psi = 0$ otherwise.

From the above, and using the definition of a reservation wage $W_1^S = \frac{T_1}{r + \delta}$, we obtain T_1 implicitly:

¹⁰The steady state measures of bachelors (B_p) and married men (M_p), together with their respective distributions of earned wages are obtained from inflow-outflow equations similar to the ones in Section 3.

$$\begin{aligned}
[r + \delta + u(\bar{p} - p_1)] \frac{T_1}{r + \delta} &= \\
&= \Psi \int_{p_1}^{\bar{p}} \left[\frac{u\lambda_0}{(\lambda n_1 + \delta)(r + \delta)} \int_{\underline{R}(p)}^{\bar{w}_p} [1 - F_p(w)] dw \right] dp + \\
&\quad + \Psi u(\bar{p} - p_1) \int_{p_1}^{\bar{p}} \frac{\underline{R}(p)}{r + \delta} dp + \\
&\quad + \int_{p_2}^{\min(p_1, \bar{p})} \left[\frac{u\lambda_0}{(\lambda n_1 + \delta)(r + \delta)} \int_{T_1}^{\bar{w}_p} [1 - F_p(w)] dw \right] dp + \\
&\quad + \int_{p_3}^{\min(p_2, \bar{p})} \left[\frac{u\lambda_0}{(\lambda n_1 + \delta)(r + \delta)} \int_{T_1}^{\bar{w}_p} [1 - F_p(w)] dw \right] dp
\end{aligned} \tag{27}$$

Women in Set 2 These women have no access to Set 1 men. As a consequence (and allowing for $\underline{p} \leq p^*$), a Set 2 woman with attractiveness y will only accept a man from the following pool:

- i) a measure N_p'' of single and employed men with $p \in [p_3, \min(p_2, \bar{p})]$ who chanced it but were unlucky and now earn $w < T_1$;
- ii) a measure $u[p_3 - \max(\underline{p}, p^*)]$ of single unemployed men with $p \in [\max(\underline{p}, p^*), p_2]$ who - while unemployed - cannot marry Set 1 women. These are men who are yet to find a job: men with $p_3 < p < p_2$ who are chancing it, or men with $p > p_2$ who set their reservation wage equal to T_1 ;
- iii) a measure $u[p'(y) - p_2]$ of single unemployed men with $p \in [p_2, p'(y)]$, with $p'(y)$ being the productivity p for which $y^*(p) = y$. These are men who - while unemployed - are once again refused by Set 1 women, and whose productivity is low enough so they accept a woman with attractiveness y .

Note that p^* , as previously defined, is obtained implicitly from $\frac{\underline{R}(p^*)}{r + \delta} = W_2^S$. The steady state measure N_p'' and the $G'(w)$ distribution of wages earned by these men are obtained from the straightforward steady state equations:

$$u\lambda_0 [F_p(w) - F_p(R(p))] = G'_p(w)N''_p (\delta + \lambda n_2)$$

and

$$u\lambda_0 [F_p(T_1) - F_p(R(p))] = N''_p (\delta + \lambda n_2).$$

Thus, we obtain:

$$G'_p(w) = \frac{F_p(w) - F_p(R(p))}{F_p(T_1) - F_p(R(p))} \quad (28)$$

and

$$N''_p = \frac{u\lambda_0 [F_p(T_1) - F_p(R(p))]}{\delta + \lambda n_2} \quad (29)$$

With $p^* < p_3$, and given the relevant combined pool of single men, the value of being single for a Set 2 woman is therefore:

$$\begin{aligned} (r + \delta)W_2^S(y) = & \int_{p_3}^{\min(p_2, \bar{p})} N''_p \int_{R(p)}^{T_1} \left[\frac{w}{r + \delta} - W_2^S \right] dG'_p(w) dp + \\ & + u[p_3 - \max(\underline{p}, p^*)] \int_{\max(\underline{p}, p^*)}^{p_3} \left[\frac{\underline{R}(p)}{r + \delta} - W_2^S \right] dp + \\ & + \Omega u [(p'(y) - p_2)] \int_{p_2}^{p'(y)} \left[\frac{\underline{R}(p)}{r + \delta} - W_2^S \right] dp, \end{aligned} \quad (30)$$

where $\Omega = 1$ for $p'(y) > p_2$ and $\Omega = 0$ otherwise.

Note that $p^* > \underline{p}$ implies that $\frac{\underline{R}(p^*)}{r + \delta} = W_2^S$, and therefore $T_2 = \underline{R}(p^*)$. On the other hand, for $p^* < \underline{p}$, we have $W_2^S > \frac{\underline{R}(p^*)}{r + \delta} (= \frac{T_2}{r + \delta})$.

4.2 Equilibrium with classes

Below, we first define an equilibrium with two classes, where the bottom class is characterised by zero marriage premium. This is essentially a summary of the best response strategies and steady-state variables derived in the previous section. Following that, we discuss the necessary existence conditions, and

show (using a numerical solution) that such an equilibrium indeed exists. We also generalise our main qualitative result to any number of classes. Finally, we provide a full characterisation of possible marriage premia patterns across male productivity types as well as marriage classes.

Given the definition of p_1, p_2, p_3 , and together with steady-state measures of various men and women, as well as earned wages, a steady-state search equilibrium with two classes is essentially a system:

$$\{T_1, T_2, \underline{R}(p), R(p), y_1, y_2, y^*(p), p^*\},$$

such that:

(i) Women's reservation matches satisfy:

$$T_1 = (r + \delta)W_1^S(w)$$

and

$$\begin{aligned} T_2 &= \underline{R}(p^*) \quad \text{for } p^* > \underline{p}, \quad \text{or} \\ T_2 &< (r + \delta)W_2^S(w) \quad \text{for } p^* < \underline{p}, \end{aligned}$$

where $W_j^S(w)$ as defined (with $j = 1, 2$).

(ii) Type p unemployed men's reservation wages and reservation matches depend on whether or not they take the marriage market into account:

(a) If they ignore the marriage market:

$$\underline{R}(p) = \frac{\lambda_0}{r + \delta} \int_{\underline{R}(p)}^{\bar{w}_p} [1 - F_p(w)] dw$$

(b) If they take the marriage market into account:

$$R(p) = T_1 \quad \text{and}$$

$$\begin{aligned} y^*(p) &= \frac{\lambda_0}{r + \delta} \int_{T_1}^{\bar{w}(p)} [1 - F(w)] dw + \frac{\lambda_0}{r + \delta} [1 - F_p(T_1)] [y_1 - y^*(p)] + \\ &\quad + \frac{\lambda n_2}{r + \delta} \int_{y^*(p)}^{y_1} [H(y_1) - H(y)] dy - \underline{R}(p) \end{aligned}$$

or

$$\begin{aligned}
R(p) &= \frac{\lambda_0}{r + \delta + \lambda H(y_1)} \int_{R(p)}^{\bar{w}(p)} [1 - F_p(w)] dw + \\
&+ \frac{\lambda_0(r + \delta)}{r + \delta + \lambda H(y_1)} [1 - F(T_1)] [V_1^S(w) - V_2^S(w)] + \\
&+ \frac{\lambda}{r + \delta + \lambda H(y_1)} \int_{\underline{y}}^{y_1} [\underline{R}(p) + y] dH(y) \quad (< T_1).
\end{aligned}$$

(iii) Employed men's reservation matches are:

$$y_1 = \frac{\lambda}{r + \delta} \int_{y_1}^{\bar{y}} [1 - H(y)] dy \quad \text{if } w \geq T_1$$

and

$$y_2 = \frac{\lambda}{r + \delta} \int_{y_2}^{y_1} [H(y_1) - H(y)] dy \quad \text{if } w < T_1.$$

(iv) Threshold male type p^* satisfies:

$$\frac{\underline{R}(p^*)}{r + \delta} = W_2^S$$

4.2.1 Existence and characterisation:

Theorem 2 *The set of parameters for which a two-class search equilibrium with zero male marriage premium in the bottom class exists is non-empty. Such an equilibrium exhibits $p^* < p_3$, and exists if $p^* < \underline{p} < p_3$ and $\underline{y} > \underline{y}'$, where $\underline{y}' = y_2(\underline{y}')$.*

Proof. All choice variables are continuous, and the equilibrium equations are twice-differentiable. To show existence, it is enough to provide a numerical solution and find a set of such equilibrium parameters (see Appendix C).

To characterise this equilibrium, first assume that $p^* > p_3$. Then, all men in Set 2 - those with $p \in (p^*, p_3)$ - will chance it, so for these men the marriage

premium is positive. Some males (those with $p < p^*$) would constitute a *third* set of men, but they would never marry.

The condition $\underline{y} > \underline{y}'$ simply says that Set 2 men are willing to marry all women not in Set 1, so there is no third female set. These women make up Set 2, which is therefore also the bottom set. Note that y_2 is positive and independent of \underline{y} when $\underline{y} < y_2$. In addition, y_2 is increasing in \underline{y} when $\underline{y} > y_2$. Hence, $\underline{y}' > 0$, and therefore it is always possible to choose a $\underline{y} \in (\underline{y}', y_1)$ such that $\underline{y} > y_2$. Consequently, set 2 men are willing to marry all women not in Set 1, which makes Set 2 the only other (bottom) female set. This also rules out a third class.

Next, assume that $\underline{p} > p_3$, so even the lowest type men chance it. But then there is always a positive marriage premium. Instead, assume $\underline{p} < p_3$. This implies $T_1 > R(\underline{p})$, and hence set 1 women reject some men. In particular, a set 1 woman will never marry an unemployed of type p_3 , or indeed any unemployed of even lower type. From the equation for T_1 it is clear that $T_1 < \bar{p} + z$ as long as λ is not infinite. Since $p_3 = T_1 - z$, we have $\bar{p} > p_3$. This ensures that there are unemployed men who chance it, and hence it rules out a one-class equilibrium. Note that $\bar{p} < p_3$ would mean that all unemployed men ignored T_1 , which would then make this T_1 sub-optimal for women. Women would set a lower reservation match, and this lower T_1 would in turn lead to the existence of additional sets of men. Since $\partial p_3 / \partial \underline{p} = 0$ for $\underline{p} < p_3$, one can always choose a low enough \underline{p} which satisfies the above condition. Together with $p^* < \underline{p}$, this ensures that the second class is in fact the bottom class, as there are only two sets of each gender. Finally, $p^* < \underline{p}$ also implies that Set 2 women marry all Set 2 men, even if these men are unemployed, meaning zero marriage premium. ■

Figure 2 below illustrates a search equilibrium with two marital classes and associated marriage premium patterns:

FIGURE 2 ABOUT HERE

For reasons of tractability, so far we have only focused on a two-class search equilibrium. However, this is without any loss of generality, either in terms of endogenous class formation¹¹ or in terms of having a zero marriage premium for men in a bottom class - that is, bottom among *any number* of classes. Indeed, we have the following result:

¹¹As shown by Burdett and Coles (1997).

Corollary 1 *Let K be the number of marriage classes and let p_K^* defined by $\underline{R}(p_K^*) = T_K$. In any search equilibrium with K classes, $p_K^* < T_K - z$ and the men in the bottom class have zero marriage premium.*

Proof. Assume K classes. Then, any equilibrium with $p_K^* < T_K - z$ is such that $MP = 0$ amongst some men in the K^{th} class, simply because \underline{p} can always be chosen such that $p^* < \underline{p} < p_3$. With two classes, $p_K^* < T_K - z$ is actually $p^* < T_1 - z = p_3$. ■

4.2.2 Marriage premia across productivities and classes

Finally, we summarise the marriage premia patterns pertaining to men of various productivities *as well as* across classes in a two-class equilibrium.

(A) Men with extremely high productivity ($p > p_1$, for whom $\underline{R}(p) > T_1$) can ignore the labour market. They qualify for male Set 1 even when unemployed. As a consequence, they only marry women with $y > y_1$ and form Class 1 partnerships with such Set 1 women. The marriage premium of these men is zero, due to the fact that labour market status is simply irrelevant for their prospects in the marriage market.

(B) Highly productive men (with $p_2 \leq p < p_1$) will not accept wages that put them in Set 2. They choose $R(p) = T_1$. If (and only if) they find such an acceptable wage, they are able to marry Set 1 women and form Class 1 partnerships with them. There is no marriage wage premium associated with marriage to Set 1 women. However, there is a marriage wage *penalty* associated with marriage to Set 2 women. Recall that although these men only accept wages that put them in Set 1, while unemployed they belong to the male Set 2, as only Set 2 women accept them. Given this, they could end up marrying a Set 2 woman who has a high enough y . Following that, however, their reservation wage goes down to $\underline{R}(p)$, which in turn generates the marriage penalty.

(C) Mid-productivity men (with $p_3 \leq p < p_2$) choose $R(p) < T_1$, and are therefore willing to accept wages that put them in male Set 2. Despite that, they can be lucky and still find a wage that puts them in male Set 1. In that case, there is a positive marriage premium associated with marrying Set 1 women and forming Class 1 couples. In contrast, marriage to a Set 2

woman leads to a marriage penalty. However, note that this penalty is lower than that of the highly productive men discussed above. This is because this earnings penalty pertains to employed men with wages between $R(p)$ and T_1 : these men find a job first, and only after that, a Set 2 female partner. This is in addition to those men who find a Set 2 partner first, subsequently drop their reservation wage to $\underline{R}(p)$ and then find a job.

(D) Low productivity men can find themselves in two situations:

(i) if $p^* < \underline{p} < p_3$, the highest available wage for those having $p \in (\underline{p}, p_3)$ is lower than T_1 . Because of this, they ignore T_1 and choose $\underline{R}(p)$ as their reservation wage. For all these men, the chosen reservation wage is higher than T_2 , therefore Set 2 women marry them irrespective of whether they are employed or unemployed. This leads to the formation of Class 2 partnerships. Importantly, there is no marriage premium associated with these men.

(ii) if $\underline{p} < p^* < p_3$, men with $p \in (p^*, p_3)$ cannot marry Set 1 women but are nonetheless accepted by Set 2 women, even while jobless. Consequently, the marriage wage premium for these men is again zero. Sadly, men with $p \in (\underline{p}, p^*)$ can never get married.

5 Empirical evidence

Theorem 1 offers clear predictions about the ranking of marriage premia across men of different productivity: less educated male cohorts should be characterised by a higher marriage wage premium. In turn, the results in Section 4 characterise the equilibrium assortative matching, with marital classes where this pattern of marriage premia is broken by a bottom class with a zero wage gap. Together, these theoretical predictions offer a good basis for a relatively straightforward empirical test.

Our primary aim in this section is therefore to estimate the marriage premium for a heterogeneous male workforce in order to recognise patterns of earnings differentials that could offer support to our model, be it with one- or two-sided heterogeneity. We use educational attainment as proxy for productivity, with higher levels of education representing higher levels of productivity. We estimate the marriage premium by regressing wages on marital status, controlling for a range of other factors, using both OLS and

fixed effects. The latter method uses the within-transformation to sweep out time-invariant individual heterogeneity, and is recommended for estimating the marriage premium - for details, see Cornwell and Rupert (1995).

In addition, we also look for patterns of classes on the marriage market that exhibit the male wage premia suggested by our theoretical model.

5.1 Data

We use two sources of data from the UK: the National Child Development Survey (NCDS) and the British Household Panel Survey (BHPS). The NCDS is a cohort study that has followed a cohort of 17,000 individuals born in a single week in 1958 in England, Scotland and Wales. Data were collected on individuals throughout their lives and we use data from waves 4, 5, 6, 7 and 8 when the individuals were aged 23, 33, 42, 46 and 50. The BHPS was collected between 1991 and 2008 and covered England and Wales, as well as parts of Scotland (later all of Scotland) and Northern Ireland. Data were collected initially on around 5,000 households, leading to around 10,000 interviews. Individuals were followed over time, and as individuals moved into new households all individuals in the new households were also interviewed. We use data from all 18 waves of the BHPS.

There are several advantages to using these two datasets. Firstly, since they are longitudinal, it is possible to control for time invariant unobservable heterogeneity by using fixed effects estimations methods. Secondly, they contain information required for the estimation of wage equations, including information on individual educational attainment. Thirdly, they provide different insights into the empirical problem. The BHPS contains individuals from different ages, meaning that for different age cohorts there may be different labour market experiences. The NCDS contains individuals of the same age, so the cohort experience will be common across all individuals. Given the difference in the characteristics of the data, using both databases can constitute a test of the robustness of the empirical results.

For our estimation we have focussed on men who are either single or are in their first marriage and who are working full time. Individuals are categorised into four groups by education status. The four groups are 1) No qualifications, 2) School leavers qualifications (qualifications obtained at the

school leaving age of 16), 3) Higher school leavers qualifications (these are qualifications that are above those for age 16 but lower than degree level, and 4) Higher education qualifications - degree and above.

The key outcome variable for this study is wages. In the NCDS individuals are asked to report their usual net pay and the period that this applies to allowing us to calculate net pay per week which we then deflate (to 1987 prices) and log. For the BHPS we construct logged, deflated (to 1987 prices) monthly income. We also include controls for age (BHPS), region of residence, social class (a marker of status based on occupation), health status¹², sector of employment and wave dummies. For the BHPS we also have full information across waves on size of employer, experience and union status. Summary statistics for the two samples are given in Table 1.

[Table 1 about here]

In Table 1 we see that individuals in the NCDS sample are more likely to be married than those in the BHPS. This may be due to the age difference between the two datasets, with the average age in the BHPS sample being around 34, whereas the NCDS sample were all 50 in 2008. Other differences include the proportion of the population with a degree, reflecting greater educational opportunities for younger age groups in the BHPS data. Also, the proportion of individuals reporting excellent health is higher in the NCDS, reflecting the different categories available when answering the self-assessed health question. Lastly we can see that in the BHPS a greater proportion work in the private sector and there is a larger share of the population grouped around higher social classes.

Tables 2, 3, 4 and 5 give the summary statistics by each of the education groups. We can see that individuals with degrees have, on average higher incomes, are in better health, less likely to be in the private sector and more likely to be of high social class. As we move down the educational groups income falls and individuals are more likely to be in a lower social class classification.

[Tables 2, 3, 4 and 5 about here]

¹²For the BHPS in wave 9 the health question was different from the other waves, we estimate models including and excluding wave 9 data and find that it makes little difference to the results.

5.2 Methods

We estimate the marriage wage premium for all individuals using fixed effects, which controls for unobservable heterogeneity and also removes all time invariant variables, such as education status, from the model. We follow Cornwell and Rupert (1995) and consider a wage equation of the form:

$$\ln(w_{it}) = \beta M_{it} + \gamma' X_{it} + \alpha_i + \varepsilon_{it},$$

where the dependent variable is the log of income, M_{it} is an indicator showing whether an individual is married, and X_{it} is a matrix of control variables.¹³ In turn, α_i captures time-invariant individual heterogeneity and ε_{it} is the standard idiosyncratic error term. The unknown parameter β is to be estimated and interpreted: if a marriage wage premium exists then β will be positive and significant.

5.3 Results

We estimated the equation above for the full sample and for each of education categories, bearing in mind the results from Theorem 1, according to which the marriage premium increases for less productive cohorts throughout the whole range, and those of Theorem 2, which predict that the increase in the marital wage gap for less productive cohorts is broken in the last class (which is characterised by a zero male marriage premium).

The results are presented in Table 6. The first row of results shows the marriage wage premium for the whole samples for the NCDS and the BHPS. The remaining rows show the results for each of the education groups.

[Table 6 about here]

Overall, we can see the existence of a marriage premium and that is highly significant, for the whole sample. When we consider the estimation by education group, and starting with the most educated men, we observe the gradient, just as predicted by both versions of our model. In both samples, the estimate is close to zero and not significant among men with degree. It

¹³For both datasets these are: region of residence, wave dummies, sector of employment, social class, health status and family size. The BHPS also includes age, size of employer, experience and union status.

becomes significant and increasing as we move to men with "below degree" and this pattern continues as we move to "school leavers".

However, there is a break in this pattern as we move to men with no qualifications, for whom the estimates are again insignificant, reflecting no marriage premium among these men. This finding suggests the existence of classes, whereby women in the lowest class and men with lowest productivity/education have no option except to marry each other - and they do so regardless of their characteristics.

In Table 7 we investigate the characteristics of men's partners in order to check if they differ across the productivity range. In particular, we focus on the characteristics of the partners of men in the lowest productivity group. Our conjecture regarding classes and the existence of no marriage wage premium for the lowest productivity group would be supported if women married to these men displayed worse characteristics, on average, than women married to men in the other productivity groups.¹⁴

For the NCDS study we have data on the partner's years of education and the partner's social class. For the latter measure we simply use the 6 point scale (where 1 is the highest social class and 6 is the lowest social class) in our regression. Since our aim is simply to highlight differences between productivity groups, it does not matter that we linearise this variable. For the BHPS we have information on a partner's highest education qualification, a partner's hourly income and a partner's parental social class. For education we focus on those with no qualifications and those with a degree (the extremes of the education qualifications) and for social class we again linearise the 6 point scale, where 1 is the highest social class and 6 is the lowest social class.

[Table 7 about here]

Examining Table 7, we can see that in both data sets, the lowest productivity men (those with no qualifications) are married to women with the worst characteristics, whichever way they are measured. Men with no qualifications are married to women with lower average years of education (in the NCDS) or women who are less likely to have a degree and more likely to have no qualification (in the BHPS). These women have, on average, lower social

¹⁴The sample sizes do not match our earlier analysis since the data for partners is not complete, leading to a number of missing cases.

class status (NCDS) or their parental social class status is significantly lower (BHPS). For the BHPS we can also see that if these women are working they earn, on average, a lower average wage than women married to men in higher productivity groups. These results provide empirical support for the existence of classes, where men in the lowest productivity group and women in the lowest class have no option but to marry each other. A consequence of that marriage pattern would be the absence of any male marriage wage premium in the bottom class, a conjecture that is in accordance with our empirical results.

6 Conclusion

The contributions of this paper fall into three categories: methodological, applied-theoretical and empirical.

In terms of method, we propose a theoretical modelling framework which looks at search behaviour in inter-linked markets. We believe the potential of such models is considerable¹⁵, as they enable the study of any scenario where agents' actions in one market depend on expectations and strategies in another market (and vice-versa), with interactions taking place in the presence of search frictions.

Potentially, there could of course be many ways of capturing such cross-market inter-dependence. One particular link, which we explore in detail, is the search behaviour of individuals in *one* market who can gain access to the *other* market only if the outcome of their search is successful enough. We call this type of search problem *constrained search*, and are able to fully characterise the resulting optimal strategy. With sequential search, it turns out that the reservation value *function* is non-monotonic.¹⁶ This reflects the trade-off between potential gains from qualifying for access to the other market, and the opportunity costs of holding out for an "entry ticket" that guarantees such access (costs that increase whenever obtaining such a "ticket" becomes increasingly unlikely). The non-monotonicity property of the best response function is quite robust in the sense that it holds as long as there are search

¹⁵The existing literature is, however, very thin. Two interesting recent papers are Kaplan and Menzio (2016) and Rupert and Wasmer (2012).

¹⁶For more details, see Bonilla et al. (2016).

frictions in the first market and utility gains from accessing the second market. Furthermore, as shown by Bonilla and Kiraly (2014), the property holds for fixed sample search problems as well.

We use this general equilibrium framework to explore the interaction between frictional labour and marriage markets, and in order to investigate the male marital premium puzzle. Following Grossbard-Shechtman and Neuman (2003), the key assumption we make is that of gender asymmetry, in the sense that men are perceived as family breadwinners - whatever the source of such beliefs. Given the non-monotonic nature of the optimal stopping strategy, the reservation wage set by single unemployed men can end up being less than the wage required to make them marriageable. Of course, this does not mean that all men are condemned to bachelorhood. Good fortune plays a role too, and some men may still land jobs with wages acceptable to women. Crucially, however, in such a search equilibrium a wedge emerges between average wages earned by married and single men respectively: the male marriage premium. In addition, our results show that ex-ante male productivity heterogeneity is neither a necessary nor a sufficient condition for the existence of such wage differentials.

The framework also lends itself to the analysis of assortative matching in the presence of search frictions, with men who differ in productivity and women who are different in terms of attractiveness. By deriving the reservation wage strategies of unemployed men, we effectively endogenise the distribution of traits (here, wages) on one side of the sorting process.¹⁷ We construct and establish the existence of an equilibrium with marital classes and show that in any such equilibrium the bottom class will always contain men with zero marriage wage premium.

Finally, we carry out the very first empirical investigation of the search theoretic explanation of marriage premium. The estimates on wage differential patterns across male productivity types and marital status, together with evidence on the existence of a bottom marriage class characterised by zero premium appear to support the rich theoretical predictions of our model.

¹⁷Strictly speaking, here this does not actually affect women's strategies, because the female reservation value is independent of the reservation wages. However, the latter is true only because the wage distributions themselves are exogenous. For more on the effect of endogenous wage distributions within this framework, see Bonilla and Kiraly (2013).

There is already a growing body of related work. Bonilla and Kiraly (2016) relax the assumption that marriages cannot be terminated and show that divorce acts like a commitment device, enabling unemployed men to marry as well. In turn, by introducing an ex-ante choice of investment in schooling, Bonilla and Kiraly (2017) contribute to the debate about why male educational attainment has fallen behind that of women. Finally, Bonilla et al. (2016) provide an explanation for the existence of the so-called "beauty premium" (and its link with marriage premium), again based on search frictions.

Taken as a whole, this research agenda suggests that labour market decisions and outcomes (such as wage differentials) may well be the result of expectations and behaviour in the marriage market, and the other way around. Search frictions exacerbate this market inter-dependence and quite possibly explain to a large extent a number of observed - and sometimes puzzling - facts. The present paper has focused on the particular question of male marriage premium, and has offered a new approach, different from the other existing (and seemingly inconclusive) explanations. More generally, however, we hope we have managed to draw attention to the elegance and potential explanatory power of models with (any) inter-linked frictional markets and constrained search.

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Table 1: Summary statistics

Variable	Mean NCDS	Std. Dev.	Mean BHPS	Std. Dev
log real weekly income	5.166	0.514		
log real monthly income			6.827	0.489
married	0.739	0.439	0.591	0.492
education - degree	0.127	0.333	0.216	0.412
education - below degree	0.353	0.478	0.341	0.474
education - school leaver (16)	0.431	0.495	0.321	0.467
education - no qualifications	0.089	0.284	0.122	0.327
age			34.511	8.558
excellent health	0.826	0.379	0.329	0.47
private sector	0.702	0.458	0.821	0.383
social class 1 (high)	0.071	0.257	0.081	0.273
social class 2	0.305	0.46	0.325	0.469
social class 3 non-manual	0.166	0.372	0.153	0.36
social class 3 manual	0.31	0.463	0.282	0.45
social class 4	0.113	0.316	0.13	0.336
social class 5 (low)	0.035	0.184	0.029	0.168
experience (days)			1714.625	2055.932
union			0.415	0.493
N	7380		19644	

Table 2: Summary statistics - Education, degree

Variable	Mean NCDS	Std. Dev.	Mean BHPS	Std. Dev
log real weekly income	5.487	0.544		
log real monthly income			7.106	0.524
married	0.748	0.434	0.545	0.498
age			34.485	8.176
excellent health	0.875	0.331	0.386	0.487
private	0.553	0.497	0.713	0.453
social class 1 (high)	0.232	0.422	0.24	0.427
social class 2	0.568	0.496	0.551	0.497
social class 3 non-manual	0.096	0.295	0.129	0.335
social class 3 manual	0.035	0.184	0.048	0.214
social class 4	0.045	0.207	0.031	0.172
social class 5 (low)	0.025	0.155	0.002	0.043
experience (days)			1188.782	1498.924
union			0.377	0.485
N	937		4251	

Table 3: Summary statistics - Education, below degree

Variable	Mean NCDS	Std. Dev.	Mean BHPS	Std. Dev
log real weekly income	5.267	0.506		
log real monthly income			6.844	0.46
married	0.742	0.438	0.592	0.492
age			34.184	8.54
excellent health	0.848	0.359	0.345	0.476
private	0.690	0.462	0.809	0.393
social class 1 (high)	0.086	0.281	0.063	0.242
social class 2	0.38	0.485	0.368	0.482
social class 3 non-manual	0.172	0.377	0.19	0.392
social class 3 manual	0.274	0.446	0.281	0.45
social class 4	0.07	0.255	0.084	0.277
social class 5 (low)	0.018	0.134	0.014	0.118
experience (days)			1657.673	1960.923
union			0.438	0.496
N	2606		6701	

Table 4: Summary statistics - Education, school leaver (16)

Variable	Mean NCDS	Std. Dev.	Mean BHPS	Std. Dev
log real weekly income	5.056	0.444		
log real monthly income			6.709	0.42
married	0.741	0.438	0.575	0.494
age			33.34	8.497
excellent health	0.808	0.394	0.297	0.457
private	0.735	0.441	0.875	0.33
social class 1 (high)	0.022	0.148	0.02	0.142
social class 2	0.208	0.406	0.212	0.409
social class 3 non-manual	0.199	0.399	0.163	0.37
social class 3 manual	0.399	0.49	0.378	0.485
social class 4	0.134	0.341	0.193	0.395
social class 5 (low)	0.037	0.189	0.033	0.178
experience (days)			1849.076	2123.723
union			0.416	0.493
N	3183		6297	

Table 5: Summary statistics - Education, no qualifications

Variable	Mean NCDS	Std. Dev.	Mean BHPS	Std. Dev
log real weekly income	4.834	0.466		
log real monthly income			6.591	0.435
married	0.702	0.458	0.714	0.452
age			38.55	8.252
excellent	0.755	0.43	0.265	0.441
private	0.797	0.403	0.906	0.292
social class 1 (high)	0.017	0.129	0.009	0.093
social class 2	0.099	0.299	0.103	0.304
social class 3 non-manual	0.086	0.28	0.066	0.248
social class 3 manual	0.414	0.493	0.446	0.497
social class 4	0.277	0.448	0.268	0.443
social class 5 (low)	0.107	0.309	0.109	0.311
experience (days)			2453.813	2641.215
union			0.419	0.493
N	654		2395	

Table 6: Marriage wage premium results

	NCDS		BHPS	
	β	N	β	N
Full sample				
Marriage wage premium	0.091*** (0.018)	7380	0.074** (0.046)	19644
Education - degree				
Marriage wage premium	0.022 (0.046)	937	-0.034 (0.074)	4251
Education - below degree				
Marriage wage premium	0.062** (0.029)	2606	0.08* (0.047)	6701
Education - school leaver				
Marriage wage premium	0.081*** (0.024)	3183	0.165* (0.091)	6297
Education - no qualifications				
Marriage wage premium	-0.011 (0.069)	654	0.042 (0.064)	2395

*, **, *** represent 10%, 5% and 1% levels of significance

All models include controls for region of residence, wave dummies, sector of employment, social class, health status and family size. BHPS models also include controls for age, size of employer, experience and union status.

Table 7: Partner's Characteristics

NCDS			BHPS			
	Partner's years of Education	Partner's Social Class	Partner has no qualifications	Partner has a degree	Partner's hourly income	Partner's parental social class
Education - degree	20.10*** (0.105)	2.403*** (0.034)	0.027*** (0.003)	0.508*** (0.011)	6.382*** (0.024)	2.765*** (0.029)
Education below degree	18.31*** (0.086)	2.886*** (0.021)	0.111*** (0.005)	0.145*** (0.006)	6.031*** (0.018)	3.542*** (0.022)
Education - school leaver	16.89*** (0.047)	3.070*** (0.019)	0.168*** (0.006)	0.079*** (0.005)	5.909*** (0.019)	3.757*** (0.021)
Education - no qualifications	16.35*** (0.113)	3.349*** (0.048)	0.377*** (0.012)	0.037*** (0.005)	5.824*** (0.025)	4.078*** (0.031)
N	5365	4646	10776	10776	8724	9602

*, **, *** represent 10%, 5% and 1% levels of significance

Appendix A: R_i ($< T$) and \hat{T}_i

For $T \in (\hat{T}_i, \bar{w}_i]$, and given $V_i^S(R_i) = J_i$, the reservation wage solves:

$$\begin{aligned} \frac{R_i}{r + \delta} &= \frac{\lambda_0}{r + \delta + \lambda_0} F(R_i) \int_{\underline{w}_i}^{R_i} \frac{R_i}{r + \delta} dF_i(w \mid w < R_i) + \\ &+ \frac{\lambda_0}{r + \delta + \lambda_0} [F_i(T) - F_i(R_i)] \int_{R_i}^T \frac{w}{r + \delta} dF_i(w \mid R_i < w < T) + \\ &+ \frac{\lambda_0}{r + \delta + \lambda_0} [1 - F_i(T)] \int_T^{\bar{w}_i} \left[\frac{w}{r + \delta} + \frac{\lambda n}{(r + \delta)(r + \delta + \lambda n)} y \right] dF_i(w \mid w > T). \end{aligned}$$

After some manipulations, this simplifies to:

$$R_i = \frac{\lambda_0}{r + \delta} \int_{R_i}^{\bar{w}_i} [w - R_i] dF_i(w) + \frac{\lambda_0 [1 - F_i(T)] \lambda n}{(r + \delta)(r + \delta + \lambda n)} y.$$

Further integration by parts leads to:

$$R_i = \frac{\lambda_0}{r + \delta} \left[\int_{R_i}^{\bar{w}_i} [1 - F_i(w)] dw + \frac{[1 - F_i(T)] \lambda n}{r + \delta + \lambda n} y \right] \quad (< T).$$

The expression for \hat{T}_i follows immediately:

$$\hat{T}_i = \frac{\lambda_0}{r + \delta} \left[\int_{\hat{T}_i}^{\bar{w}_i} [1 - F_i(w)] + \frac{[1 - F(\hat{T}_i)] \lambda n}{r + \delta + \lambda n} y \right] \quad (> \underline{R}_i).$$

With no search frictions in the labour market ($\lambda_0 \rightarrow \infty$), clearly $R_i(T) = \bar{w}_i$. Note that \hat{T}_i is a function of y . Since $\partial F_i(\hat{T}_i)/\partial \hat{T}_i > 0$, it is true that $\partial \hat{T}_i/\partial y > 0$. One can solve for y to get:

$$y = \frac{(r + \delta + \lambda n) \left[(r + \delta) \hat{T}_i - \lambda_0 \int_{\hat{T}_i}^{\bar{w}_i} [1 - F_i(w)] dw \right]}{\lambda_0 \lambda n [1 - F_i(\hat{T}_i)]}.$$

Then, $\lim_{\hat{T}_i \rightarrow \bar{w}} y = \infty$ and, since \hat{T}_i is an invertible function, $\lim_{y \rightarrow \infty} \hat{T}_i = \bar{w}_i$.

Appendix B: Other equilibria

Assume $\hat{T}_H > \bar{w}_L$. With $\frac{\partial T}{\partial x} > 0$ and $\underline{R}_H < \hat{T}_L < \bar{w}_L < \hat{T}_H$, it follows that $\underline{x} < \hat{x}_L < \bar{x} < \hat{x}_H$.

Now assume $\hat{T}_H < \bar{w}_L$. For $x \in [\underline{x}, \hat{x}_H]$, there is always a unique equilibrium where only employed men can get married and both types do. The range of possible equilibria and the necessary and sufficient conditions in terms of x are as follows:

- (a) An equilibrium with $MP_L > MP_H > 0$ obtains iff $x > \hat{x}_H$.
- (b) An equilibrium with $MP_L > 0$ and $MP_H > 0$ obtains iff $\bar{x} < x < \hat{x}_H$.
- (c) An equilibrium with $MP_L > 0$ and $MP_H = 0$ obtains iff $\hat{x}_L < x < \bar{x}$.
- (d) An equilibrium with $MP_i = 0$ ($i = H, L$) obtains iff $\underline{x} < x < \hat{x}_L$.

Appendix C: Proof of Theorem 2 (numerical solution)

As all choice variables are continuous, it is enough to show that a two-class equilibrium with the desirable properties exists for a given set of parameters. Consider the following values for the relevant parameters of the model: $\lambda_0 = 1.5$; $\delta = 0.1$; $r = 0.1$; $\lambda = 1$; $\bar{p} = 200$; $z = 200$; $u = 0.1$; $\underline{y} = 0.4$; $\bar{y} = 1$. Given a range of productivities $[\underline{p}, \bar{p}]$ where each productivity has the same measure u of unemployed, for men also we are dealing with an uniform distribution of types. Finally, let the distribution of wages be uniform as well, with support on $[p, p + z]$ for each type p .

With two classes (by construction), we have: $n_1 = [1 - H(y_1)]$ and $n_2 = H(y_1)$. Also, note that women's arrival rate is very sensitive to changes in u and the productivity range.

Equilibrium:

We pick the equilibrium that is most difficult to construct, i.e. the one with $p_2 > \bar{p}$. Since p_2 is endogenous, showing that such an equilibrium exists is not trivial precisely because of this constraint on \bar{p} .

First, note that with the above inequality constraint, some of the indicator functions in our value functions will be zero. Second, we need $y_2 < y$. With our numerical values, this inequality is indeed satisfied. Next, we compute y_1 :

$$y_1 = 0.615619$$

Now we can compute T_1 , given that $p_3 \equiv T_1 - z$. We obtain:

$$T_1 = 327.08844$$

This is clearly lower than the highest wage in the market, which is $\bar{p} + z = 400$.

Next, we compute p_2 and then show that it is indeed higher than \bar{p} . Making use of the equation for $R(p)$ and equating it to T , we get $\hat{T}(p)$. Following that, solve $\hat{T}(p) = T_1$ to obtain:

$$p_2 = 258.944839 \quad (> \bar{p})$$

Recall that p_3 was defined by $T_1 - z$, so $p_3 = 127.08844$. This allows us to compute the value of being single for a Set 2 woman (W_2^S), as a function of (any) p . We can then compare this with $\frac{R(p)}{r+\delta}$ (the value of being married to a p man). Clearly, we need the latter to be greater than the former. The explicit solution to W_2^S is:

$$W_2^S = \frac{\int_{p_3}^{\bar{p}} \left[N_p'' \int_{R(p)}^{T_1} \frac{w}{r+\delta} dG_p'(w) \right] dp + u \int_{\underline{p}}^{\bar{p}} \frac{R(p)}{r+\delta} dp}{r + \delta + u(\bar{p} - \underline{p}) + \int_{p_3}^{\bar{p}} N_p'' [G_p'(T_1) - G_p'(R(p))] dp},$$

where $G_p'(T_1) - G_p'(R(p)) = 1 - 0 = 1$.

We find that $W_2^S < \frac{R(p)}{r+\delta}$ holds for $p > -61.786371 (= p^*)$. This means that there is no lower (third) class if $p > -61.786371$. All men with p between \underline{p} and p_3 constitute a subset of Set 2 - the subset of men with zero marriage premium. These men are born in Set 2.

Overall, a non-empty set (p^*, p_3) , together with the condition on y ensure the existence of our two-class equilibrium for a broad range of parameters.

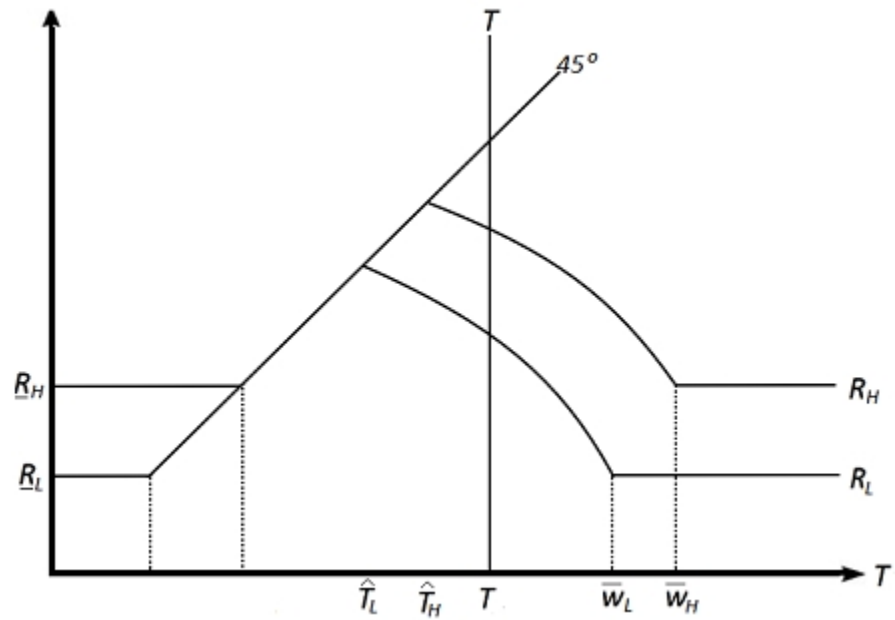


Figure 1

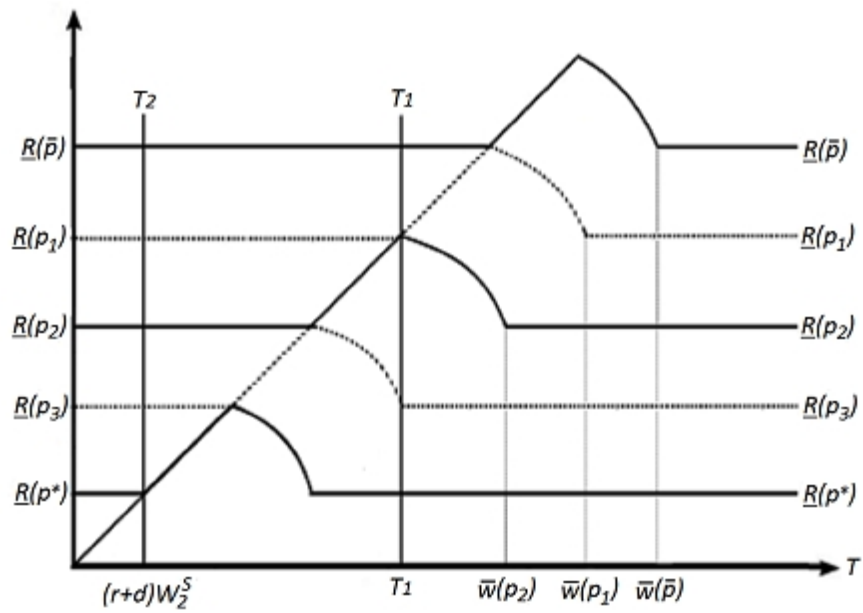


Figure 2