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## Structural filters for monetary analysis: the inflationary movements of money in the euro area

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EUROPEAN CENTRAL BANK

**WORKING PAPER SERIES**

**NO. 470 / APRIL 2005**

**STRUCTURAL FILTERS  
FOR MONETARY ANALYSIS**

**THE INFLATIONARY  
MOVEMENTS OF MONEY  
IN THE EURO AREA**

by Annick Bruggeman,  
Gonzalo Camba-Méndez,  
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EUROPEAN CENTRAL BANK



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### STRUCTURAL FILTERS FOR MONETARY ANALYSIS

### THE INFLATIONARY MOVEMENTS OF MONEY IN THE EURO AREA <sup>1</sup>

by Annick Bruggeman <sup>2</sup>,  
Gonzalo Camba-Méndez <sup>3</sup>,  
Björn Fischer <sup>3</sup>  
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# CONTENTS

Abstract	4
Non-technical summary	5
1 Introduction	7
2 The link between money growth and inflation	9
3 Structural filters for euro area M3	11
3.1 Doubled-sided filter	11
3.2 One-sided filter	13
3.3 Empirical results	13
4 Real-time use of structural filters	16
5 Conclusions	17
References	19
Figures	21
Table	30
European Central Bank working paper series	31

## Abstract

The quantity theory of money predicts a positive relationship between monetary growth and inflation over long-run horizons. However, in the short-run, transitory shocks to either money or inflation can obscure the inflationary signal stemming from money. The spectral analysis of time series provides filtering tools for removing fluctuations associated with certain frequency movements. However, use of these techniques in isolation is often criticised as being an oversimplistic statistical exercise potentially void of economic content. The objective of this paper is to develop 'structural' filtering techniques that rely on the use of spectral analysis in combination with a structural economic model with well identified shocks. A 'money augmented' Phillips curve that links inflation to money tightness and demand shocks of medium to long-term persistence is presented. It is shown that medium to long-term movements in inflation are mostly associated with the estimated monetary indicators.

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## Non-technical Summary

The quantity theory of money predicts a positive relationship between monetary growth and inflation over long-run horizons. However, in the short-run, transitory shocks to either money or inflation can obscure the inflationary signal stemming from money. From a monetary policy perspective it is important to identify those movements in M3 that are associated with inflationary pressures, and discard other movements associated with transitory non-inflationary developments. Communication from monetary authorities often makes references to developments in inflation and money over ‘medium-run’ or ‘long-run’ horizons. Furthermore, monetary authorities’ mandates are very often defined over these same horizons. This could be taken as an indication that some isolation of the long and medium-run movements of economic time series is implicitly made both by those responsible for setting the mandate and those responsible for the implementation of monetary policy. Spectral analysis techniques provide filtering tools for removing those fluctuations in economic time series associated with certain frequencies, i.e. short, medium or long-run. However, use of these filtering tools in isolation is often criticised as being an over-simplistic statistical exercise potentially void of economic content. The objective of this paper is to develop structural filtering techniques that rely on the use of spectral analysis on well identified structural economic models.

We make use of a money augmented Phillips curve to explain fluctuations in inflation. This may be thought as corresponding to a two pillar view of inflation fluctuations. Namely, that long-run developments in inflation are explained by long-run movements in money, and short to medium term developments by the output gap. The techniques we develop in this paper do not prejudge what frequency-movements in money should be isolated to estimate inflation, but rather provide those movements in money as a by-product of the modelling strategy. Furthermore, we identify those movements in inflation associated with demand shocks and movements associated with money shocks, i.e. monetary looseness or tightness, where these are well identified orthogonal shocks.

Our framework enables us to study separately what drives long-run movements in inflation from factors behind short-run movements. Therefore, it formulates in a quantitative manner a definition for medium to long-run developments, and hence it provides valid quantitative analysis to back communication from policy makers on long-run economic developments. The money series identified by the structural filters developed in this paper provide valid indicators of monetary tightness. They represent that component of money that feeds into inflation movements with a periodicity higher than 9 months.

This paper applies this new structural filter to euro area data for the low inflation period

1986Q1 to 2003Q2. The average rate of growth of euro area M3 over the period 1994 to 2003 was 5.2% with a minimum value of  $2\frac{1}{2}\%$  on the second quarter of 1995 and a maximum of  $8\frac{1}{2}\%$  on the second quarter of 2003. We report measures of monetary tightness exclusively linked with inflation developments, and show that large movements in money over this period were not linked to inflationary pressures. These series of money linked to inflationary pressures displayed a much smoother pattern than the actual series of M3, fluctuating between values of  $4\frac{1}{2}\%$  and  $6\frac{1}{2}\%$  over this same period. This money tightness indicator is also shown to be the key driving force behind fluctuations in inflation.



# 1 Introduction

The quantity theory of money predicts a positive relationship between monetary growth and inflation over long-run horizons. However, in the short-run, transitory shocks to either money or inflation can obscure the inflationary signal stemming from money. From a monetary policy perspective it is important to identify those movements in M3 that are associated with inflationary pressures, and discard other movements associated with transitory non-inflationary developments. Communication from monetary authorities often makes references to developments in inflation and money over ‘medium-run’ or ‘long-run’ horizons. Furthermore, monetary authorities’ mandates are very often defined over these same horizons. This could be taken as an indication that some isolation of the long and medium-run movements of economic time series is implicitly made both by those responsible for setting the mandate and those responsible for the implementation of monetary policy. Spectral analysis techniques provide filtering tools for removing those fluctuations in economic time series associated with certain frequencies, i.e. short, medium or long-run. However, use of these filtering tools in isolation is often criticised as being an over-simplistic statistical exercise potentially void of economic content. The objective of this paper is to develop structural filtering techniques that rely on the use of spectral analysis on well identified structural economic models. Among other things, this will enable us to identify that component of inflation associated with monetary shocks and medium to long-run persistency.

We follow Gerlach (2003) in using a money augmented Phillips curve to explain fluctuations in inflation. Gerlach (2003) regards this as a valid interpretation of the ECB’s monetary policy strategy, and in his view this corresponds to a two pillar view of inflation fluctuations; namely, that long-run developments in inflation are explained by long-run movements in money, and short to medium term developments by the output gap. Gerlach (2003) used a pre-filtered series of money, where use was made of a low-pass filter to isolate long-run movements. Nelson (2003) also highlighted that studying the relationship between the long-run movements in inflation and long-run movements in money is particularly valuable for central banks because their aim is setting the steady-state rate of inflation. The techniques we develop in this paper do not prejudge what frequency-movements in money should be isolated to estimate inflation, but rather provide those movements in money as a by-product of the modelling strategy. Furthermore, we identify those movements in inflation associated with demand shocks and movements associated with money shocks, i.e. monetary looseness or tightness, where these are well identified orthogonal shocks.

The framework for monetary analysis currently used by the ECB already provides an approach to derive the signal from monetary growth associated with risks to future price

stability; see Masuch, Pill, and Willeke (2001). This framework relies on the combination of information obtained from various econometric models with the detailed institutional analysis conducted by experts. This has enabled a thorough understanding of monetary trends in the euro area and the early identification of shocks affecting the information content of monetary data. For example, the close monitoring of monetary data has allowed the timely detection of statistical outliers, e.g. the strong monthly monetary growth in January 1999. This outlier was partly related to the introduction of a new statistical reporting scheme at the start of Stage Three of Economic and Monetary Union (EMU). The analysis has also helped to identify shocks to monetary growth that were related to economic factors not likely to be inflationary. For example, the strong monetary growth between mid-2001 and end 2003 was to a large extent related to the increased liquidity preference of investors in an environment of high economic and financial market uncertainty. In addition, in the context of the economic analysis, shocks to inflation have also been detected that were temporary and unrelated to monetary growth, e.g. oil or food price shocks.

This type of monetary analysis has been highly valuable since the start of Stage Three of EMU. However, there seems to be a case for complementing it with the novel modelling strategy presented in this paper. Our framework enables us to study separately what drives long-run movements in inflation from factors behind short-run movements. Therefore, it formulates in a quantitative manner a definition for medium to long-run developments, and hence it provides valid quantitative analysis to back communication from policy makers on long-run economic developments. The money series identified by the structural filters developed in this paper provide valid indicators of monetary tightness. They represent that component of money that feeds into inflation movements with a periodicity higher than 9 months. It is our view that these indicators could be a useful complement to the analysis of the developments in annual M3 growth in the euro area currently conducted.

Isolation of long-run movements relies on forecasts at the end of the sample. It is thus important to study the real time performance of our methods to assess its usefulness for a central bank. Results presented below will show that extending the original money growth series with forecasts stemming from a money demand model reduces the size of the revisions to our monetary tightness indicators.

The remainder of the paper is organised as follows. Section 2 presents a review of the relevant empirical literature on the long-run link between monetary growth and inflation. Section 3 derives structural filters for monetary analysis. Section 4 discusses issues linked with the real-time use of the proposed filters. Finally, Section 5 concludes.

## 2 The link between money growth and inflation

The quantity theory of money states that there exists a long-run relationship between monetary growth and inflation that can be formulated as follows:

$$\Delta p = \Delta m - \Delta y + \Delta v$$

where  $\Delta p$  denotes inflation,  $\Delta m$  nominal monetary growth,  $\Delta y$  real output growth and  $\Delta v$  changes in the income velocity of money. If  $\Delta v$  and  $\Delta y$  are constant, a ‘one-to-one’ long-run relationship between nominal monetary growth and inflation exists. If only  $\Delta v$  is constant a ‘positive’ long-run relationship between nominal monetary growth and inflation exists.

There is an extensive empirical literature documenting the medium to long-term link between money growth and inflation. Lucas (1980) showed that using transformations that give less weight to the low-frequency components of M1 growth and inflation in the US results in a very weak relationship between these transformed or ‘filtered’ series. However, giving more weight to the low-frequency components results in an almost proportional relationship, as predicted by the quantity theory of money. Recent examples of this approach for the euro area can be found in Gerlach (2004), Neumann (2003) and Jaeger (2003). Gerlach and Neumann formulate a ‘money augmented’ Phillips curve model in which the expected inflation rate, or low-frequency component of inflation is approximated by the low-frequency component of the rate of growth of M3 per unit of output, i.e. M3 divided by real GDP. Results in Neumann (2003) show that the rate of growth of this money measure has been the dominant source of inflation in the euro area over the period 1986-2002. Gerlach (2004) also finds that the low-frequency component of inflation is closely related to the rate of growth of M3 per unit of output both over the period 1980-2002 and 1980-1990. Furthermore, Gerlach (2004) concludes that a model in which inflation is explained by lagged trend money growth and a lagged output gap measure fits better when estimated on data covering the low inflation period following 1991 than when estimated over the 1970s and 80s. This may be associated with the more forward-looking nature of current inflation developments.

Gerlach and Svensson (2003) and Vega and Trecroci (2002) included the real money gap as a substitute for or a complement to the output gap in a Phillips curve setting. Both studies concluded that the real money gap and the output gap both help to explain inflation developments in the euro area.

Jaeger (2003) conducted a frequency domain analysis of the link between money and inflation. He showed that the coherence (which can be interpreted as the correlation between two frequency components) between inflation and money growth is very high at the lower frequencies but much less so for frequencies similar to those associated with business cycle fluctuations, i.e. between 2 and 8 years.

Studies conducted with structural vector autoregressive (S-VAR) models also find evidence on a positive link between inflation and money. However, these studies are not conclusive about the appropriate horizon over which money is more correlated with inflation. For instance, Shapiro and Watson (1988) conclude that shocks to money play an important role in explaining the variability of inflation (over 80%) in the US, even at short time horizons. However, results of Christiano, Eichenbaum, and Evans (1999) for the US suggest that the impact of a shock to M2 on prices is clearer at longer horizons. Also the results of Altig, Christiano, Eichenbaum, and Linde (2002) for the US suggest that a shock to M2 only impacts upon prices a year and a half later. A similar result for the euro area is reported in Sousa and Zaghini (2004) who use the same identification strategy of Peersman and Smets (2001). The forecast error decompositions of prices in both models show that shocks to euro area M3 do not explain much of the forecast error variance in prices at relatively short horizons. However, the cumulative effect of the shocks to M3 to the variability in prices increases significantly with the horizon.

Other studies relied on data for a large group of countries over a relatively long time span. The main problem common to most of the panel data studies is that they disregard the fact that the behaviour of velocity and real output growth may not be similar across countries. One way to handle this problem is to rely on a panel of pooled time series and cross section data and to use a fixed effects specification, which allows the constant of the regressions to differ across countries. In this context, de Grauwe and Polan (2001) find that these fixed effects tend to increase more than proportionally with the growth rates of money, suggesting that the long-run developments in velocity and real output do indeed differ across countries. At the same time, allowing for these fixed effects does not alter the authors' conclusion that the link between monetary growth and inflation is much stronger for high inflation countries (with an average inflation rate above 20%). Nelson (2003), however, suggests that De Grauwe and Polan's rejection of the long-run link between monetary growth and inflation for low inflation countries is unjustified. His main critique is that they do not allow inflation to respond to monetary growth with a lag. Nelson shows that allowing for dynamics in a satisfactory manner leads to results that are "decidedly more favourable to the quantity theory".

A major critique that can be voiced on all these studies is that they do not give any indication on the out-of-sample forecasting performance. However, Nicoletti-Altimari (2001) replicates the analysis of Stock and Watson (1999) for the euro area and concludes on the basis of the out-of-sample forecasting performance that M3 growth is a valid indicator of future inflation in the euro area. After comparing the out-of-sample forecasting performance of a whole range of monetary, financial and real variables he comes to the conclusion that

quarterly M3 growth provides valuable information on developments in inflation over the medium to long term, i.e. between 9 and 12 quarters ahead, whereas this is not the case for shorter lags. The interpretation of these results is however somewhat complex concerning the link between money growth and inflation, as the forecasting horizon and the smoothness of the signal are indistinguishably linked in this exercise. However, it is likely that both factors contribute to this result. First, monetary growth seems to lead inflation by around 1 to  $1\frac{1}{2}$  years. Second, monetary growth seems to have information content for the longer-term inflation trend rather than for short-term fluctuations in inflation. This might explain why for short horizons monetary growth does not play a role, as short-term shocks to inflation, that are mostly unrelated to short-term shocks in monetary growth, contribute importantly to the overall variance of inflation over that horizon. By extending the forecasting horizon, one automatically smoothes inflation, thereby increasing the signal to noise ratio.

### 3 Structural Filters for euro area M3

We use a money augmented Phillips curve as the underlying model to describe fluctuations in inflation.

$$\pi_t = \Phi_m(L) \Delta m_t + \Phi_y(L) gap_t + \varepsilon_t$$

where  $\pi_t$  denotes the inflation rate,  $\Delta m_t$  denotes money growth,  $gap_t$  an output gap indicator,  $\varepsilon_t$  an iid shock, and where  $\Phi_m(L)$  and  $\Phi_y(L)$  are polynomial operators of leads and lags. The structure of these polynomial operators will depend on the type of filter we choose. Two choices will be made, namely a double-sided filter and a one-sided filter. In the technical sections below  $\mathbf{y}_t$  should be understood as  $\pi_t$  and  $\mathbf{x}_t$  corresponds with the vector series  $(gap_t, \Delta m_t)'$ . Note also that  $gap_t$  is an economic variable which by construction displays a zero mean. Therefore, the model above satisfies long run homogeneity restrictions, namely that the steady state inflation rate does not depend on the output gap, but exclusively on money growth.

#### 3.1 Double-sided filter

The modelling framework presented in this section is an extension of that in Brillinger (1981, ch. 8). Let  $\{\mathbf{y}_t\}_{t=1}^{\infty}$  and  $\{\mathbf{x}_t\}_{t=1}^{\infty}$  be  $m \times 1$  and  $n \times 1$  vector-valued second order stationary series with mean  $\mathbf{c}_y$  and  $\mathbf{c}_x$  respectively and cross spectral density matrices given by  $\mathbf{h}_{yy}(\lambda)$ ,  $\mathbf{h}_{xx}(\lambda)$  and  $\mathbf{h}_{xy}(\lambda)$  for  $0 \leq (\lambda) \leq 2\pi$ . The problem to be addressed is that of searching for the  $m \times n$  filter  $\{\mathbf{b}_u\}_{u=-\infty}^{\infty}$  such that  $\mathbf{y}_t^*$  is near  $\mathbf{y}_t$  where:

$$\mathbf{y}_t^* = \boldsymbol{\mu} + \sum_{k=-\infty}^{\infty} \mathbf{b}_{t-k} \mathbf{x}_k \quad (1)$$



The minimum of  $E \{\boldsymbol{\varepsilon}'_t \boldsymbol{\varepsilon}_t\}$  for  $\boldsymbol{\varepsilon}_t = (\mathbf{y}_t - \mathbf{y}_t^*)$  is achieved for:

$$\begin{aligned}\boldsymbol{\mu} &= \mathbf{c}_y - \left( \sum_{u=-\infty}^{\infty} \mathbf{b}_u \right) \mathbf{c}_y \\ \mathbf{b}_u &= \boldsymbol{\psi}(\mathbf{B}(\alpha), u) = \frac{1}{2\pi} \int_0^{2\pi} \mathbf{B}(\alpha) e^{iu\alpha} d\alpha\end{aligned}\quad (2)$$

where the transfer function  $\mathbf{B}(\lambda) = \mathbf{h}_{yx}(\lambda)\mathbf{h}_{xx}^{-1}(\lambda)$ . The function  $\boldsymbol{\psi}(\mathbf{B}(\alpha), u)$  is simply the inverse Fourier transform associated with the transfer function  $\mathbf{B}(\lambda)$ . Model (1) provides a representation for the movements of  $\mathbf{y}_t$  as a function of  $\mathbf{x}_t$  and  $\boldsymbol{\varepsilon}_t$ , where  $\boldsymbol{\varepsilon}_t$  is orthogonal to the other inputs by construction.

**Identification of shocks.** The inputs  $\mathbf{x}_t$  are not orthogonal, but in certain instances it might be possible to identify orthogonal input shocks by incorporating further restrictions possibly based on economic theory in the manner which is standard in the structural VAR literature. In what follows we will use the denotation  $\overline{\mathbf{A}}$ , to denote the transpose conjugate of a complex matrix  $\mathbf{A}$ . Identification of the shocks requires to identify a structure such that  $\mathbf{h}_{xx}(\lambda) = \mathbf{P}(\lambda)\boldsymbol{\Psi}(\lambda)\overline{\mathbf{P}(\lambda)}$ , and where the spectral density matrix  $\boldsymbol{\Psi}(\lambda)$  is diagonal, i.e. coherencies are always zero. This would allow to rewrite the decomposition of the variance of  $\mathbf{y}_t$  as:

$$\boldsymbol{\Sigma}_{yy} = \int_0^{2\pi} \mathbf{h}_{yy}(\lambda) d\lambda = \sum_{k=1}^n \int_0^{2\pi} \mathbf{B}_j(\lambda)\overline{\mathbf{B}_j(\lambda)} + \int_0^{2\pi} \mathbf{h}_{\varepsilon\varepsilon}(\lambda) d\lambda \quad (3)$$

where  $\mathbf{B}_j(\lambda)$  denotes the  $j$ -th column of  $\mathbf{B}(\lambda)$  and  $\mathbf{B}(\lambda) = \mathbf{B}(\lambda)\mathbf{P}(\lambda)\boldsymbol{\Psi}(\lambda)^{\frac{1}{2}} = \mathbf{B}(\lambda)\mathcal{P}(\lambda)$ . Retrieving the signal from these identified shocks is easily accomplished by using an appropriate decomposition of the transfer function. The transfer function to extract the signal of the  $k$ -th identified shocks is defined as

$$\mathbf{O}^k(\lambda) = \mathbf{B}(\lambda)\mathcal{P}_k(\lambda) (\mathcal{P}(\lambda)^{-1})_{k,\cdot} \quad (4)$$

where  $\mathcal{P}_k(\lambda)$  denotes the  $k$ -th column of  $\mathcal{P}(\lambda)$ , and  $(\mathcal{P}(\lambda)^{-1})_{k,\cdot}$  denotes the  $k$ -th row of  $(\mathcal{P}(\lambda))^{-1}$ . Computation of the filters  $\{\boldsymbol{o}_t^k\}$  corresponding to this transfer function is accomplished in the usual manner, i.e.  $\boldsymbol{o}_u^k = \boldsymbol{\psi}(\mathbf{O}^k(\alpha), u)$ .

**Band pass filtering.** It is easy to show that the minimum of  $E \{\boldsymbol{\varepsilon}'_t \boldsymbol{\varepsilon}_t\}$  for components of a certain frequency is achieved for  $\tilde{\mathbf{b}}_u = \boldsymbol{\psi}(\mathbf{F}(\alpha)\mathbf{B}(\alpha), u)$  where  $\mathbf{B}(\alpha)$  is defined as above and  $\mathbf{F}(\alpha)$  for  $0 \leq \alpha \leq 2\pi$  is the transfer function of a band-pass filter, i.e. for given values for  $\alpha_0$  and  $\mathcal{K}$  it is given by:

$$\mathbf{F}(\alpha) = \begin{cases} \mathbf{I} & \text{if } |\alpha \pm \alpha_0| \leq \mathcal{K} \\ \mathbf{0} & \text{otherwise} \end{cases}$$

or what is the same the integration to compute  $\{\mathbf{b}_t\}$  in (2) is run over the relevant frequencies.<sup>1</sup> This allows us to isolate that component in money exclusively related with fluctuations in inflation with a periodicity larger than 9 months. In our application below, this filter would be simply defined as  $\tilde{\sigma}_u^k = \psi(\mathbf{F}(\alpha)\mathbf{O}^k(\alpha), u)$ , for values of  $\alpha_0 = 0$  and  $\mathcal{K} = 1.99$ .

### 3.2 One-sided filters

Let  $\{\mathbf{y}_t\}_{t=1}^{\infty}$  and  $\{\mathbf{x}_t\}_{t=1}^{\infty}$  be  $m \times 1$  and  $n \times 1$  vector-valued second order stationary series. We estimate the model

$$\boldsymbol{\alpha}(L)\mathbf{y}_t = \boldsymbol{\delta}(L)\mathbf{x}_t + \boldsymbol{\varepsilon}_t \quad (5)$$

where  $\boldsymbol{\alpha}(L)$  and  $\boldsymbol{\delta}(L)$  are standard polynomial lag operators. The implicit one-sided filter in the model above is given by  $\boldsymbol{\alpha}(L)^{-1}\boldsymbol{\delta}(L)\mathbf{x}_t$ . The roots in the polynomial lag  $\boldsymbol{\alpha}(L)^{-1}\boldsymbol{\delta}(L)$  determine what kind of filter will be applied to the right hand side variables. If the roots are positive the weights are exponentially decaying. This filter can be understood then as a Cogley (2002) type of filter. If additionally, one of the roots is close to one, this translates into substantial smoothing (the weight of the filter for distant observations is still significant), whereas a value close to 0 leads to little smoothing (the weight of the filter goes to 0 very quickly).

### 3.3 Empirical Results

**The Data.** Use is made of euro area quarterly data over the period 1986Q1 to 2003Q2. The choice of this relatively short period was based on a number of considerations. First, the period of the 70s and early 80s was characterised by high and volatile inflation and may therefore not be so informative about the current situation of low and stable inflation rates, see Figure 1. This is suggested by de Grauwe and Polan (2001) and Gerlach (2003) who conclude that the link between monetary growth and inflation is only valid for countries or periods with high inflation. By concentrating on a sample period with low inflation rates this paper therefore takes a cautious approach. Without explicitly addressing the issue of the possible different regimes, we avoid that our results could be solely driven by the ‘high inflation part’ of the sample considered. At the same time, this implies that we run the risk of introducing a bias against the role of money. Second, standard unit root and stationarity tests suggest that within this shorter sample period inflation can be treated as a stationary process, while this is not the case when the sample period starts in 1970 or 1980. Indeed, the null hypothesis that inflation for the above-mentioned period is stationary cannot be rejected

<sup>1</sup>It is also easy to show that this solution is equivalent to the solution obtained by a two step method in which the series  $\mathbf{y}_t$  and  $\mathbf{x}_t$  are both prefiltered with the same band-pass filter and use as inputs in the model.

at the 5% significance level with the Kwiatkowski-Phillips-Schmidt-Shin test. Furthermore, the null hypothesis of a unit root can be rejected at the 5% significance level using the Phillips-Perron test. The output gap has been computed using the Hodrick and Prescott (1997) filter with  $\lambda = 1600$ .

**Identification of Shocks for the double-sided filter.** In order to identify a double-sided structural filter of monetary tightness it simply remains to formulate a structure for the spectral density matrix  $\mathbf{h}_{xx}(\lambda)$ , and where, as indicated above,  $\mathbf{x}_t = (\text{gap}_t, \Delta m_t)'$ . We pursue the following strategy. First, we derive the Wold representation  $\mathbf{x}_t = \mathbf{C}(L)\boldsymbol{\varepsilon}_t$ , where  $\boldsymbol{\varepsilon}_t$  is a  $2 \times 1$  iid process with mean zero and covariance matrix  $\boldsymbol{\Sigma} = \mathbf{P}\mathbf{P}'$ . The spectral density matrix of  $\mathbf{x}_t$  will then be written as  $\mathbf{h}_{xx}(\lambda) = \mathbf{P}(\lambda)\overline{\mathbf{P}(\lambda)}$ ; where  $\mathbf{P}(\lambda) = \mathbf{C}(e^{-i\lambda})\mathbf{P}$ . For this definition of  $\mathbf{P}(\lambda)$ , computation of the transfer function of the structural filter follows from equation (4). Second, in determining the structural shocks, i.e. in identifying matrix  $\mathbf{P}$ , we follow a procedure similar to the one used in the structural VAR literature. The specification implies that both the output gap and the rate of growth of M3 are explained by demand and money ‘tightness’ shocks. We impose the necessary restrictions for identification using a Choleski decomposition, thereby obtaining monetary shocks which are orthogonal to demand shocks. Thus, the identification strategy assumes that demand shocks respond only with a lag to money shocks. By contrast, monetary authorities can respond to contemporary demand shocks by tightening or loosening monetary conditions. In addition, money growth responds contemporaneously to demand shocks because liquidity requirements fluctuate with the state of the economy.

**Results.** Figure 2 shows the multiple coherence of money growth and the output gap with respect to inflation for the time period 1986Q1 to 2003Q2. This measure can be understood in a simplified way as the multiple correlation between these two series and inflation at different frequencies or cycle lengths. Figure 2 reveals that the multiple coherence is very high for the low and business cycle frequencies. Figure 2 also plots the coherence between inflation and the output gap once the impact of money growth is removed from both these series, i.e. the ‘partial’ coherence. It reveals that the output gap provides additional information to explain inflation with respect to money growth for movements associated with business cycle frequencies. However, it provides little additional information to explain the very long-run movements of inflation.

A first issue that has to be solved in the double-sided approach is how to find an appropriate length of the window for which the filter is defined. In this paper the choice of the appropriate length of the window is based on the results of the Akaike Information Criterion.



The results suggest that a window length of 8 quarters is desirable. A second issue relates to the end-point problem. Since the equation for inflation also includes eight quarters of leading money growth, monetary data have been extended up to 2005Q2, using forecasts, see section 4 below for more details.

Figure 3 shows the filter weights for the annual rate of growth of money stemming from the applied double-sided structural filter. As stated above it is useful for policy to isolate the medium to long run movements. We have thus implemented a structural double-sided filter that identifies money shocks of a persistency larger than 9 months. Figure 4 shows the resulting smoothed series of monetary growth. The double-sided filter smooths money growth considerably. For example, the strong portfolio shifts affecting M3 growth in the 1993-1996 episode, which had no impact on inflation, are basically completely smoothed out by the filter. The decomposition of inflation into the contributions of money and demand shocks as described in equation (3) is displayed in Figure 5. It clearly illustrates that most of the movements in inflation are associated with money shocks.

For the estimation of the one-sided filter, the general to specific approach of Hendry and Krolzig (2001) was used to arrive to an optimal lag structure, allowing for lags up to order seven. The resulting equation of the modified Phillips curve including money growth is shown below (standard errors in brackets):

$$\pi_t = \underset{(0.1)}{0.62}\pi_{t-3} + \underset{(0.1)}{0.21}gap_{t-3} - \underset{(0.1)}{0.22}gap_{t-6} + \underset{(0.04)}{0.07}\Delta m_{t-3} + \underset{(0.04)}{0.09}\Delta m_{t-7} + \varepsilon_t$$

where  $\pi_t$  is the annualized quarter-on-quarter inflation rate,  $gap_t$  is the output gap and  $\Delta m$  represents the annualized quarter-on-quarter money growth. All coefficients are significantly different from zero. Lag 7 for money is significant, in line with the expectation that money is leading inflation by around  $1\frac{1}{2}$  years. The filter weights corresponding to the Phillips curve type of equations with M3 growth are shown in Figure 6. The weights show a relatively slow decay over time, e.g. money contributes mainly via long-term movements to inflation. This is further illustrated in Figure 7, that shows the implied smoothed series of the annual growth rate of money, mean and phase shift corrected.<sup>2</sup>

The decomposition of inflation into the contributions of money growth, the output gap and the unexplained part can be derived following the principles outlined in subsection 3.2. Figure 8 shows that money growth is the main contributor to the smooth underlying trend of inflation.

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<sup>2</sup>The phase shift correction was performed using cross-correlation analysis between the original money growth series and the filtered series. In addition, the annualised quarter-on-quarter rates of growth were transformed into annual rates of growth.

## 4 Real-time use of structural filters

The previous section suggests that smoothed series of monetary growth may serve as an additional tool for monitoring those monetary developments that may lead to inflation. However, such a tool is only useful for policy purposes, if several conditions are fulfilled in real time. First, given the time lag of the transmission of a monetary policy shock to inflation, it is important that the well known leading signal of money growth to inflation, see for example Nicoletti-Altimari (2001), is preserved. Second, revisions of the filters over time should not lead to major distortions in detecting the relevant signal, especially around turning points. This has several consequences for the filtered series for money growth, both for the one-sided and the double sided filter. In order to extract the signal from current monetary developments for future inflation, the considerable phase shift induced for the money signal stemming from the one-sided filter of the augmented Phillips curve has to be counterbalanced with a considerable reliance on forecasts. The double-sided filter is additionally dependent on such forecasts because the weighting scheme relies not only on past and contemporaneous values but also on future values. Therefore, when new data become available, and forecasts are partially replaced by actual figures, the filtered series are revised.

We assess the ‘real time’ performance of the filters by studying the size of the revisions caused by a) the replacement of forecast values with observed values and b) the recursive estimation of the filter weights. This definition is therefore only partial, as it ignores revisions of the underlying raw data, which can also distort the results.

In order to construct the structural filters in real time, we compare two different scenarios for the computation of the eight quarters of forecasts:

1. Use of ARIMA models.<sup>3</sup>
2. Use of a structural model. Money growth is forecasted from the money demand model of Calza, Gerdesmeier, and Levy (2001), and where use is made of an ARIMA(1,1,1) for forecasting real GDP, and an ARIMA(0,2,1) for forecasting the GDP deflator. Additionally, interest rates and oil prices are assumed to be constant during the forecast horizon.

In order to reproduce the real-time procedure, the models are estimated recursively and forecasts are made each time for the next eight quarters. First, they are estimated for the period 1986Q1-1991Q4 and used to predict M3 over the following 8 quarters. Then they are re-estimated up to 1992Q1 and another 8 quarters of forecasts are produced, and so on and

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<sup>3</sup>An ARIMA(0,1,1) model was used to forecast money growth; the characteristics of the other ARIMA models are described in the second scenario.

so forth until 2001Q2. The money demand equation from 2001Q2 onwards is fixed at the 2001Q2 estimates to avoid short-run instability problems in the money demand model.

Due to the small sample size and the relatively low number of turning points, a formal analysis of revisions and especially of the speed in detecting turning points in real time is not possible. However, one can present the main stylised facts in a more heuristic way. In order to have a visual assessment of the impact of the two different strategies on the real-time analysis, Figure 9 shows filtered M3 growth in real time for the double-sided as well as the one sided filter. Some very clear stylised facts can be detected for both filters. First, revisions, especially around turning points, are substantial for the first variant of using only ARIMA models for extending the series. Second, those revisions are substantially reduced, especially around turning points, when the series of monetary growth is extended with forecast from the money demand model.

Our analysis revealed that revisions start to be insignificant after ten quarters, so that we considered the filter to be final when the end point of money growth was  $t + 10$ . Table 1 shows the size of the mean squared revisions  $t$  to  $t + 10$  for the time period 1992Q1 to 2001Q4, e.g. a measurement of the size of the revision from the information available in real time to the final information. It is clear from the table that relying on forecasts from a money demand model reduces the relative mean squared revision considerably for both filters.

## 5 Conclusions

This paper developed structural filtering techniques that relied on the use of spectral analysis on well identified structural economic models. This enabled us to identify that component of inflation associated with monetary movements which have a medium to long-run persistency. In deriving the filters we control for the influence of the output gap. It is our view that our structural filters could serve as valid quantitative analysis to back communication from policy makers on long-run monetary developments.

We applied these filters to euro area data for the low inflation period 1986Q1 to 2003Q2. The average rate of growth of euro area M3 over the period 1994 to 2003 was 5.2% with a minimum value of  $2\frac{1}{2}\%$  on the second quarter of 1995 and a maximum of  $8\frac{1}{2}\%$  on the second quarter of 2003. We have shown that large movements in money over this period were not linked to inflationary pressures and reported measures of monetary tightness exclusively linked with inflation developments. These series displayed a much smoother pattern, fluctuating between values of  $4\frac{1}{2}\%$  and  $6\frac{1}{2}\%$  over this same period. This money tightness indicator was also shown to be the key driving force behind fluctuations in inflation. The analysis

conducted in this paper is founded on a money augmented Phillips curve and some simple identification assumptions. To the extent that this represents an oversimplistic view of the economy, our results may be subject to criticism. Use of the modelling strategy presented in this paper with a more elaborate model of the economy is currently pursued by the authors.

The one-sided as well as the double-sided structural filters for money rely on forecasts at the current end of the time series. In that respect, we concluded that using structural forecasting models substantially reduced the size of the revisions of our money tightness indicator, especially around turning points, and provide policy relevant information.

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Figure 1: Euro Area Monetary Growth and Inflation (in percent).

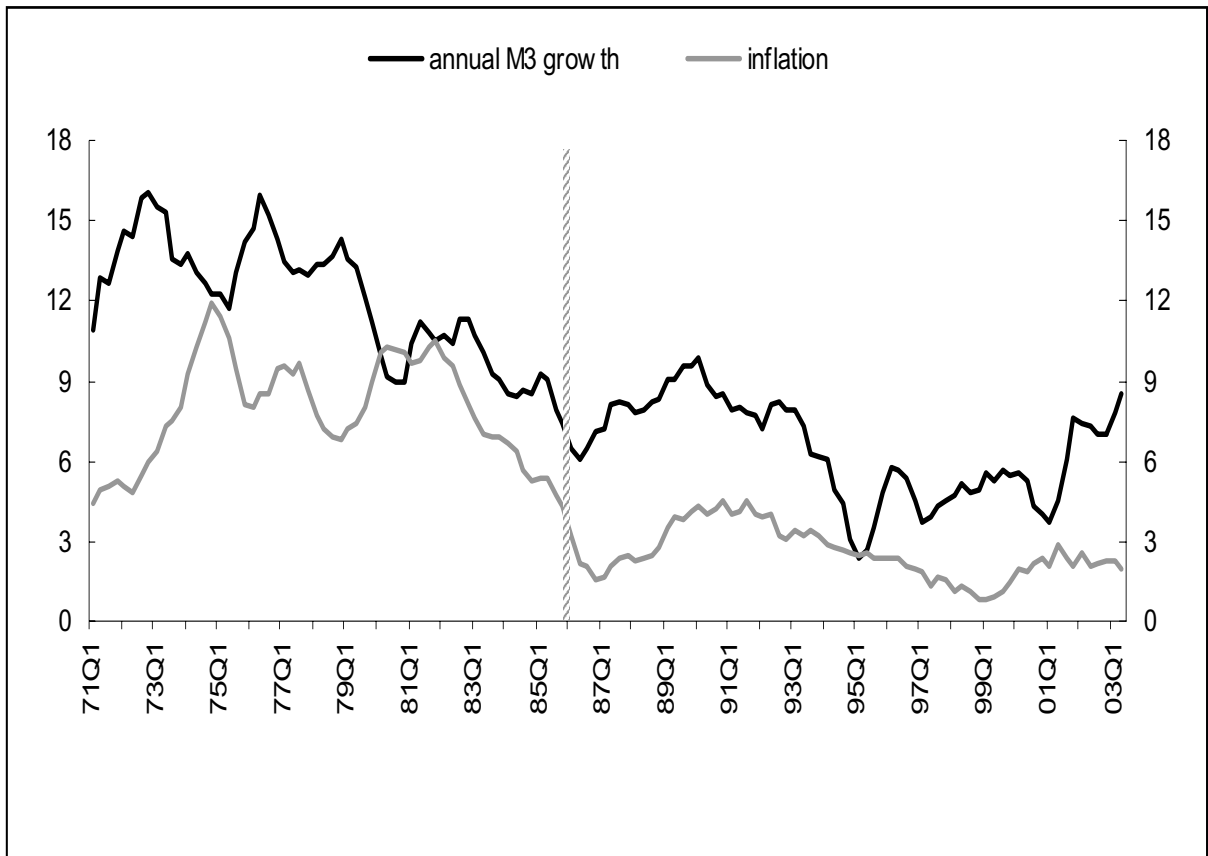


Figure 2: Multiple Coherence between Inflation and M3 Growth and the Output Gap.

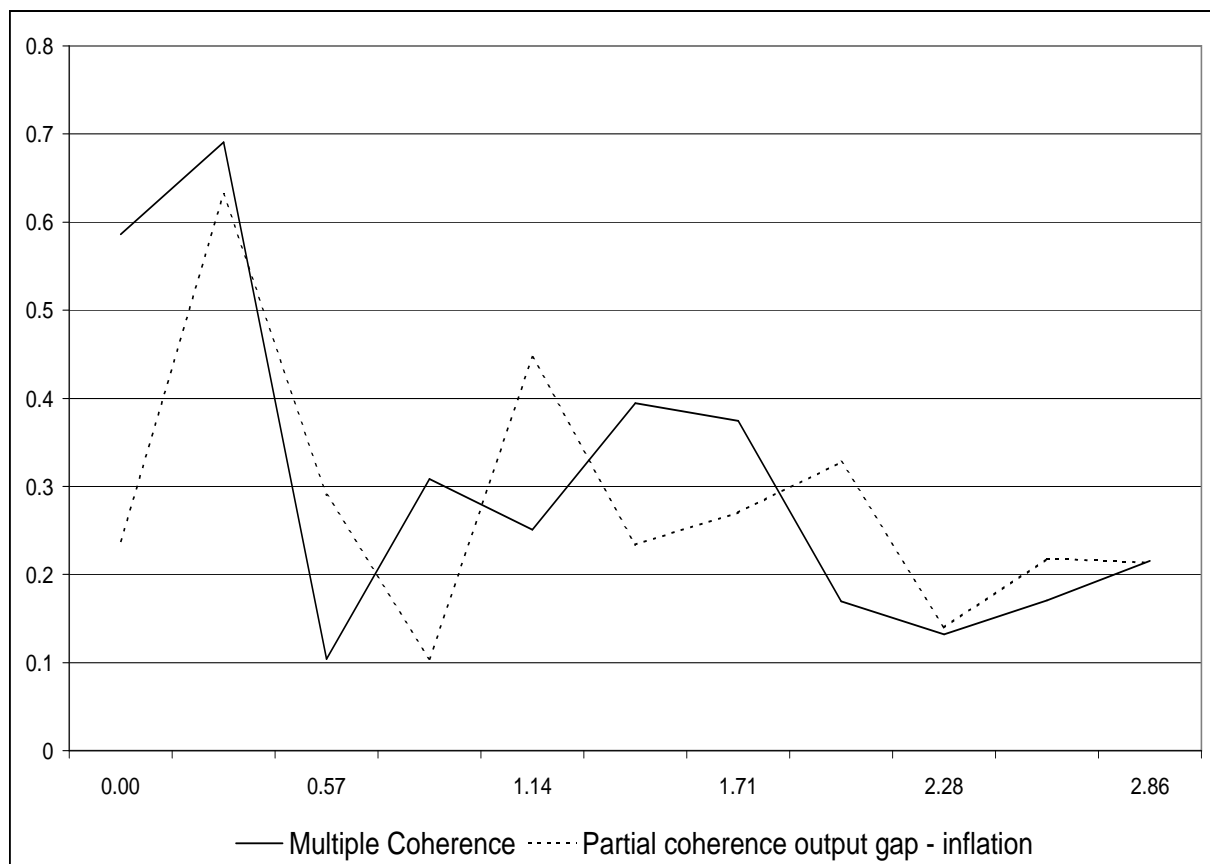




Figure 3: Weights for Double-sided Monetary Tightness Filter. (> 9 months).

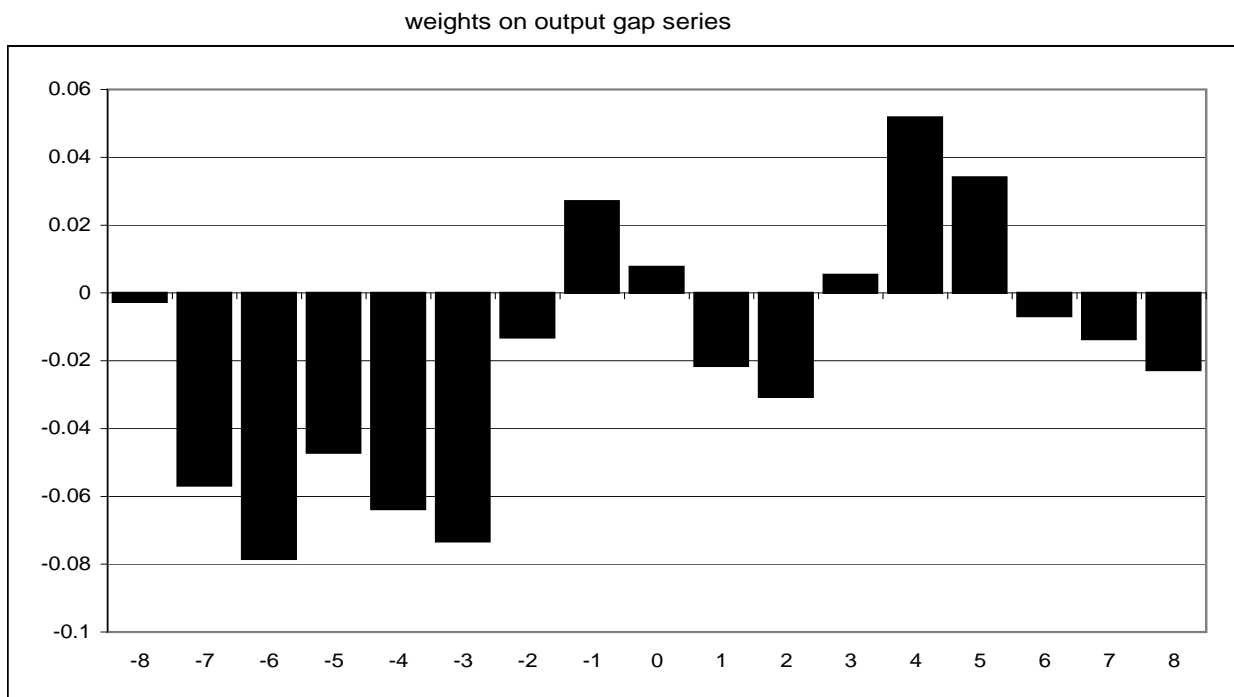
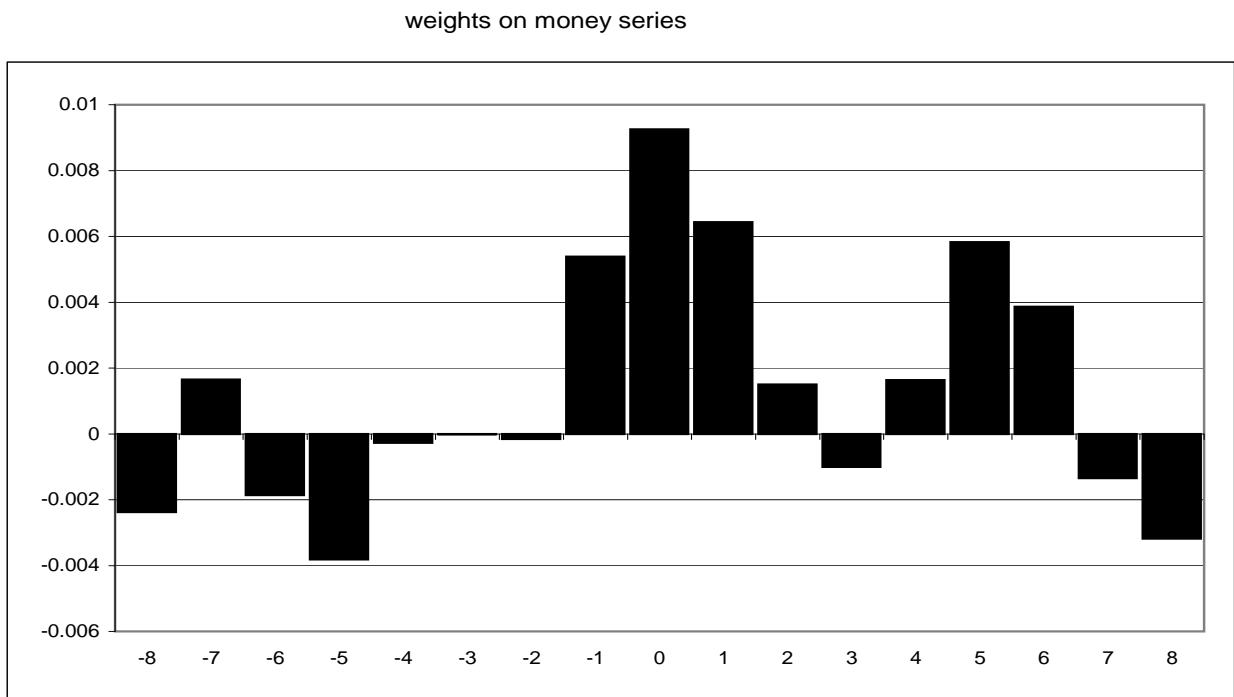


Figure 4: Money and Inflationary (> 9 months) Money Movements. Double-sided.

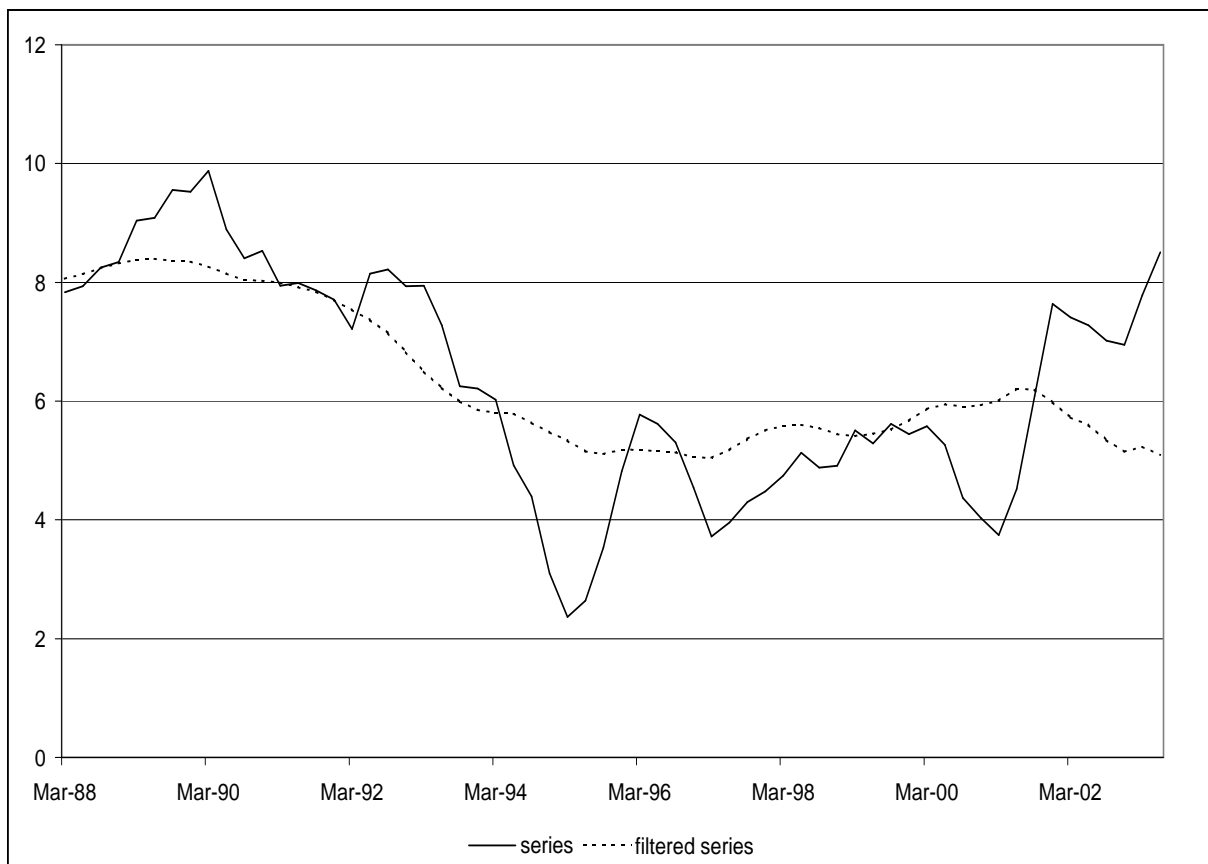


Figure 5: Decomposition of Shocks to Inflation (>9 Months Movements). Double-Sided.

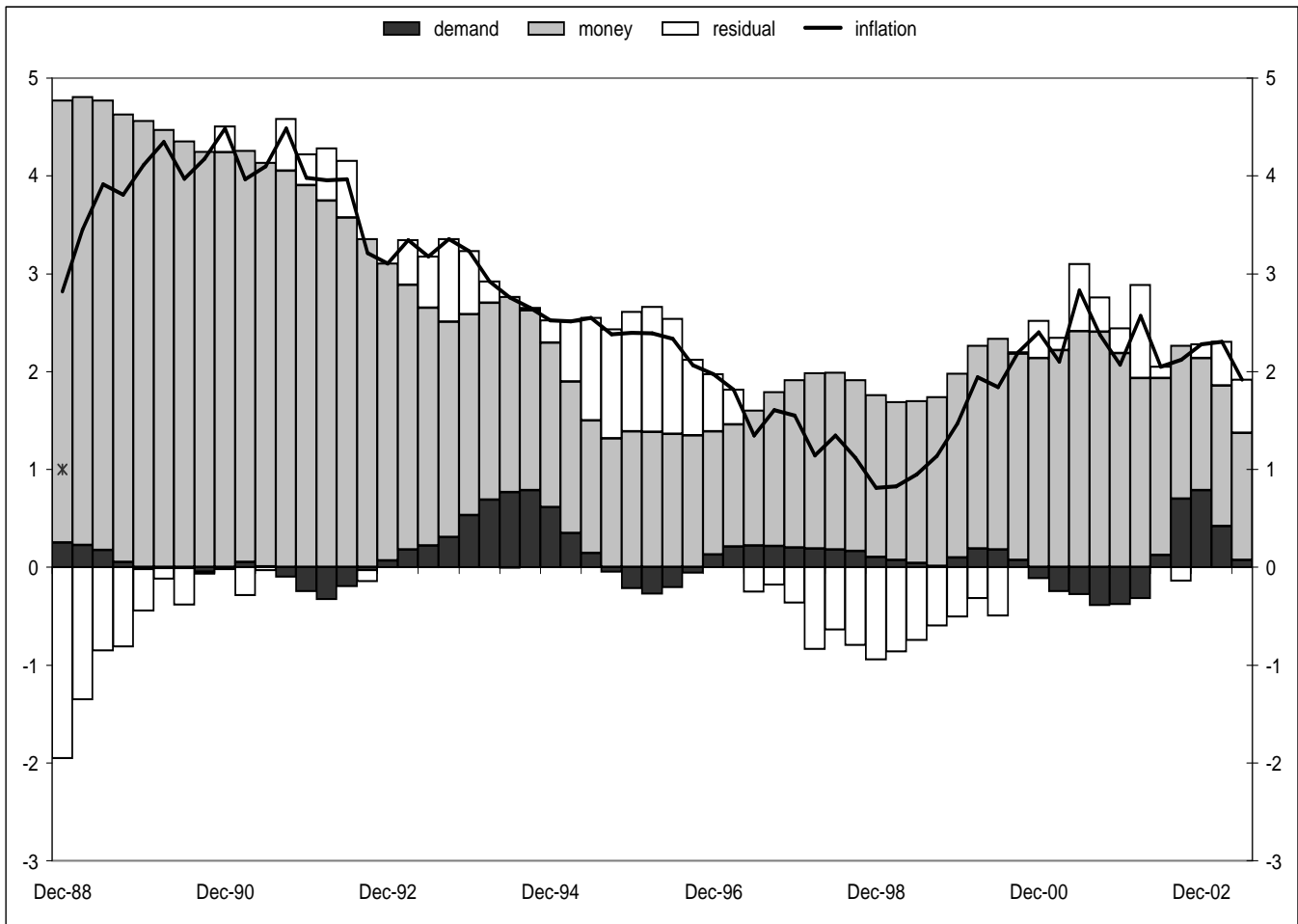


Figure 6: Weights for One-sided Monetary Tightness Filter.

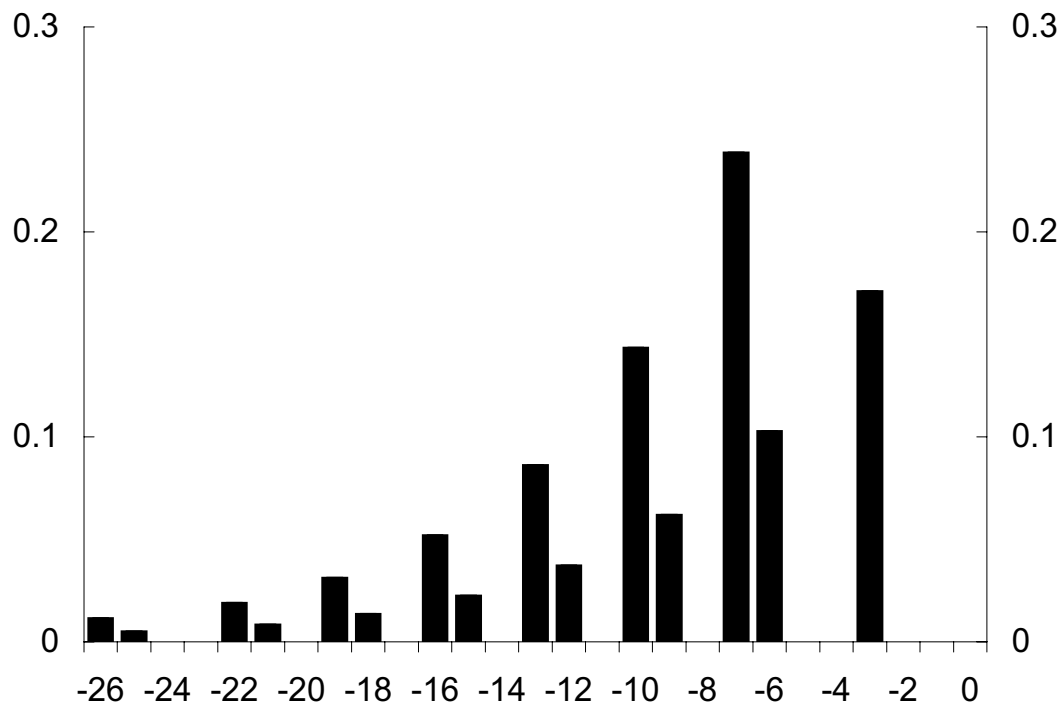


Figure 7: Money and Inflationary Money Movements. One-sided.

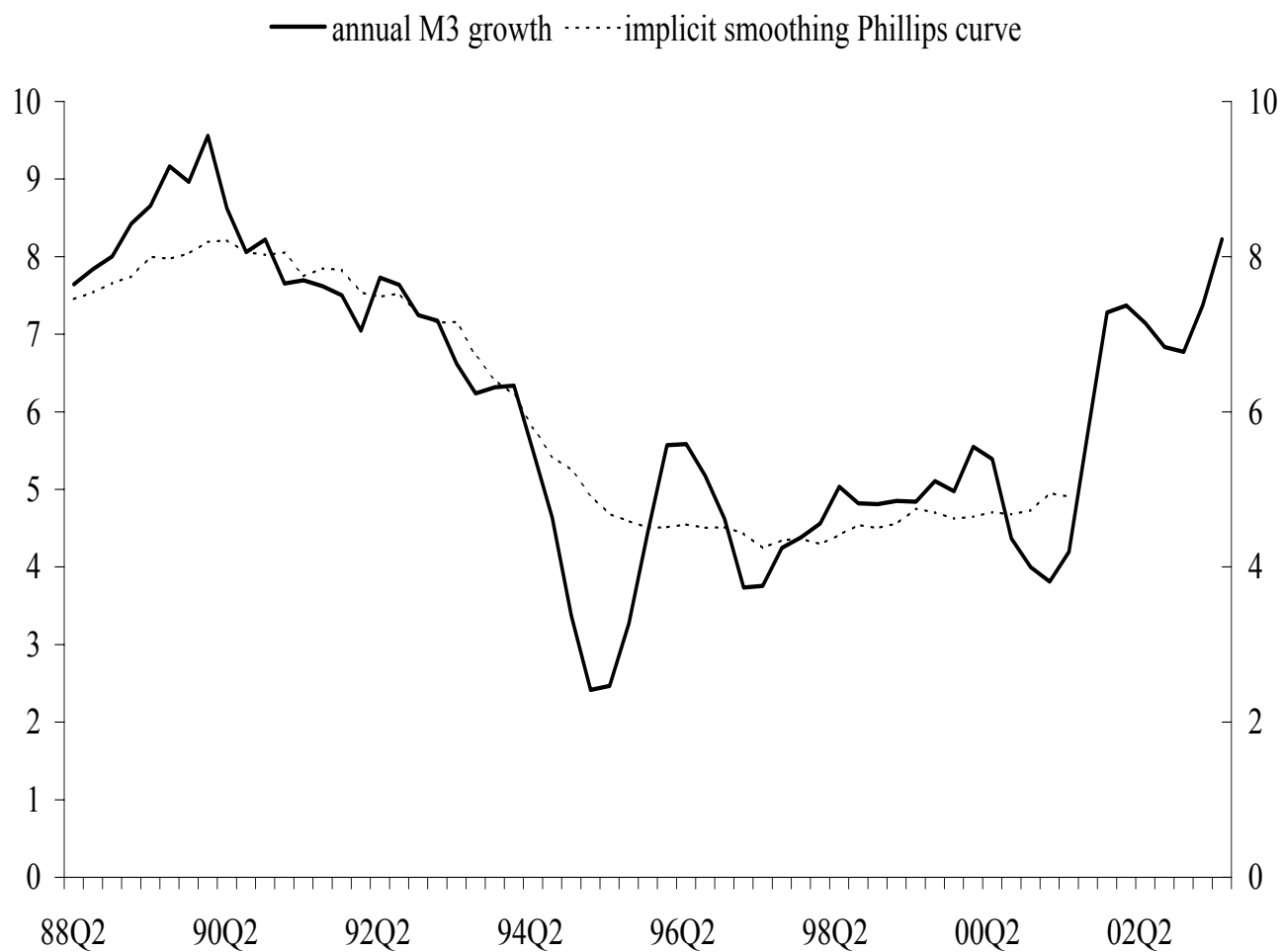


Figure 8: Decomposition of Shocks to Inflation. One-Sided.

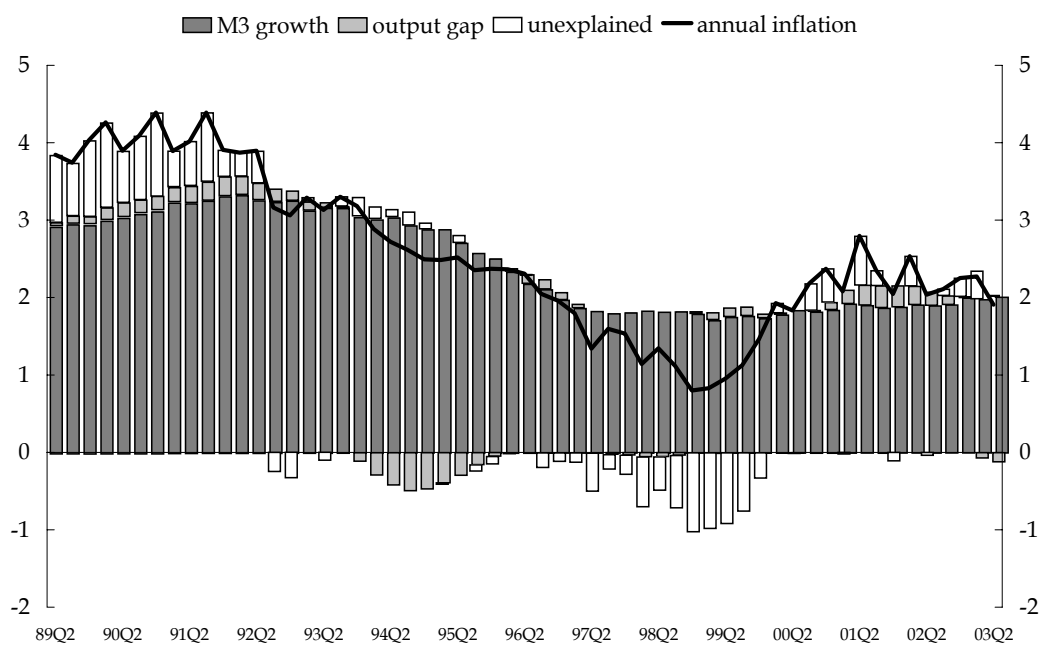


Figure 9: Real Time Estimation of Monetary Tightness Filters.

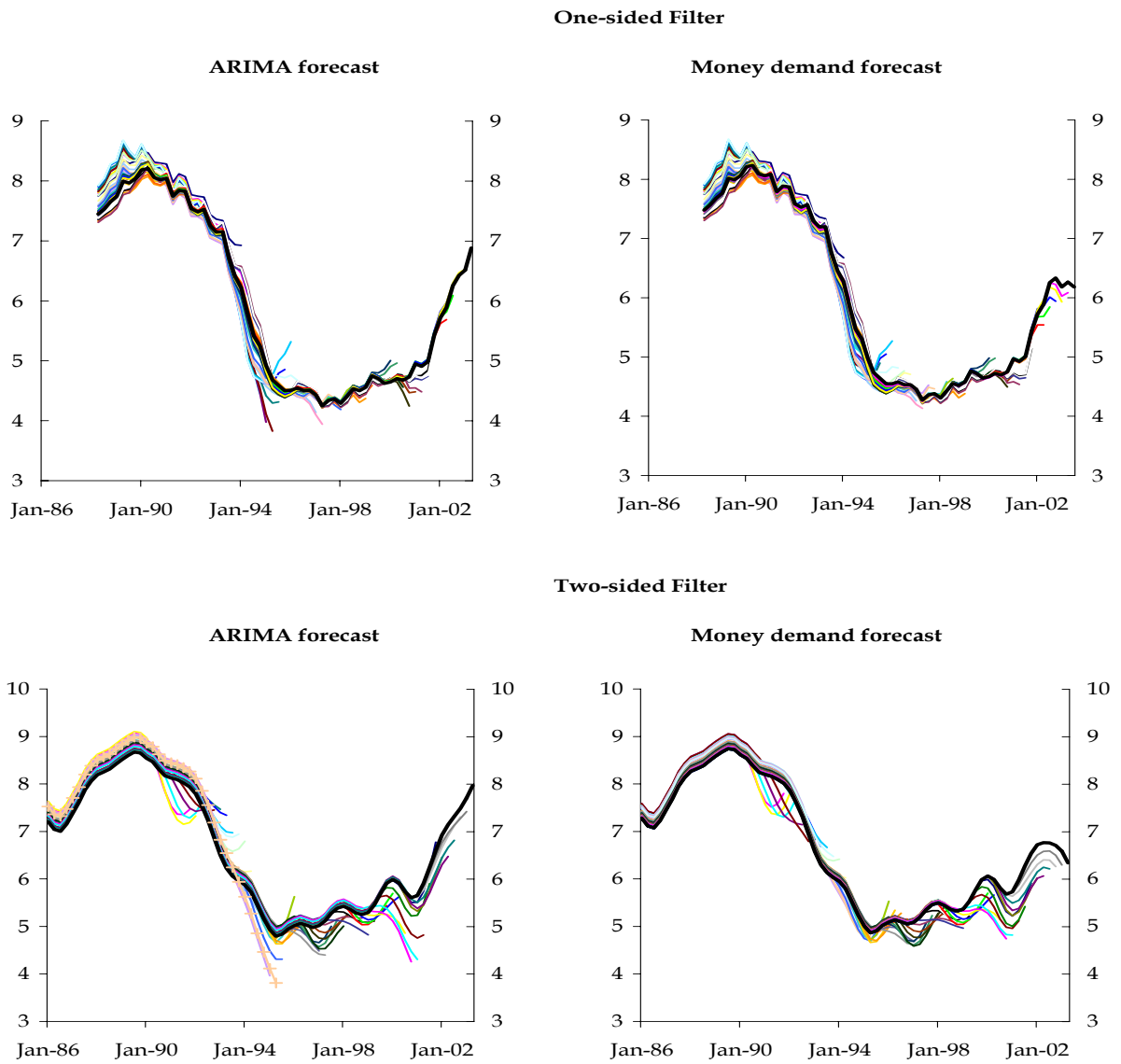


Table 1: RMSE of Real-Time Estimated Monetary Tightness Filters.

	One sided Filter	Two Sided Filter
ARIMA Forecast	0.09	0.12
Money Demand Forecast	0.05	0.03



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