

Almeida, Rita K.; Faria, Marta

Article

The wage returns to on-the-job training: Evidence from matched employer-employee data

IZA Journal of Labor & Development

Provided in Cooperation with:

IZA – Institute of Labor Economics

Suggested Citation: Almeida, Rita K.; Faria, Marta (2014) : The wage returns to on-the-job training: Evidence from matched employer-employee data, IZA Journal of Labor & Development, ISSN 2193-9020, Springer, Heidelberg, Vol. 3, pp. 1-33, <https://doi.org/10.1186/2193-9020-3-19>

This Version is available at:

<https://hdl.handle.net/10419/152394>

Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

Terms of use:

Documents in EconStor may be saved and copied for your personal and scholarly purposes.

You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.

If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.



<http://creativecommons.org/licenses/by/2.0/>

ORIGINAL ARTICLE

Open Access

The wage returns to on-the-job training: evidence from matched employer-employee data

Rita K Almeida¹ and Marta Faria^{2*}

* Correspondence:

martalincefaria@gmail.com

²Católica Lisbon, SBE, Lisbon, Portugal

Full list of author information is available at the end of the article

Abstract

Skills shortages and skill mismatch are a pressing concern for policymakers in several developing countries, and in East Asia specifically. Providing on-the-job training can be an effective policy tool to shape the skills of the existent workforce to the specific needs of the firms. This paper explores a unique data set of matched employer-employee data for Malaysia and Thailand to estimate the wage return to on-the-job training in these two countries. Exploring propensity score matching estimates, we show that the average wage returns to on-the-job training are 7.7% for Malaysia and 4.5% for Thailand. Furthermore, we find evidence that the wage returns to on-the-job training are higher for males than for females in Malaysia and that, for both countries, returns are higher for workers with at least secondary education.

JEL Classification codes: J24; J30

Keywords: On-the-job training; Wages; Matched employer-employee data

1. Introduction

Many economists have emphasized the importance of human capital accumulation for growth (e.g. Lucas, 1988; Romer, 1990; Aghion-Howitt, 1998). Human capital accumulation is done throughout life, but more than one half of this lifetime accumulation is done on-the-job after completing formal schooling (e.g., Heckman et al., 1998). In spite of the importance, much more is known about the investment in formal schooling, and specifically on the returns to formal schooling, than on the investment in on-the-job training and their returns.

This paper explores a matched employer-employee data set with unique information on formal firms and their workers for two developing countries with very different levels of development, Malaysia and Thailand¹. According to the World Economic Outlook Database from 2005, the per capita GDP in 2002 was 3,880 USD in Malaysia and 1,994 USD in Thailand and in 2004 was 4,624 USD and 2,521 USD, respectively². Between 2002 and 2004, these economies were growing at similar rates (5.6% and 6.1%). In terms of population, Thailand more than doubles the size of Malaysia (65m vs. 25m). Both countries present a high literacy rate of around 93%. East Asian countries have been drawing a great deal of attention to themselves. Impressive growth rates, competitive wages and high levels of education of the workforce are just some of the reasons why it has now become so interesting to study these countries. Having

a detailed data-set from these two countries offers the possibility of studying the dynamics of these labor markets and their investment in job training.

In developing countries, governments are increasingly concerned with the rapidly changing demand for skills and the slow response of the general and vocational schooling tracks to adjust the provision of skills. As a consequence, many employers complain about the lack of skills and education of their workforce. Policymakers are thus increasingly concerned that the supply of skills in the labor market does not keep pace with the demand. The investment by firms in on-the-job training is one important way to mitigate this skills' gap as it develops job relevant skills among the existing workforce.

The evidence on both the incidence and the economic returns to on-the-job training is generally scarce in developing countries³. And it is unclear how different the returns should be in developed and developing countries. On the one hand, the returns to the investment in job training (as well as in schooling) could be higher in developing than in developed countries simply because skilled labor is scarcer in developing countries (e.g. Psacharopoulos and Patrinos, 2004). On the other hand, if skilled labor and capital are complements, the returns to this investment could be smaller in developing countries, where capital is relatively scarce.

In theory, whether workers with and without on-the-job training receive, all else constant, significantly different wages will also relate to whether the training offered general or firm specific skills. It may also relate to whether there are differences in the competitiveness of the local labor markets. When the labor market is perfectly competitive and training is general, workers will support the cost of job training through lower wages during that same period. Once training is received, the worker will be paid the equivalent to his marginal productivity, which we now assume to be higher (e.g. Becker, 1964). But when training is firm specific the costs and benefits will likely be shared between the firm and the worker depending on the bargaining power of each one of them. In principle, the worker will receive a lower wage at the time of the training, to account for his share of the costs, and a higher wage after the training event, depending on the benefit he could extract from the firm (e.g. Leuven and Oosterbeek, 2001). If the labor market is not competitive and firms are able to pay a wage lower than the worker's marginal productivity, firms will only want to invest in training if the increase in productivity is higher than the effect in the growth rate of wages (e.g. Acemoglu and Pischke, 1999). Even in this scenario, there is no theoretical reason for the wages to decrease after the training program. They should increase or remain constant. In sum, no matter the assumptions we have, theory predicts that after participating in a training event the worker's wage should increase or stay invariable. Finally, as training is a decision variable for the firm one expects wage returns to job training to be a lower bound estimate for the impact of training in firm productivity⁴.

This paper estimates whether the firm's investment in job training translates into higher wages for the workers in Malaysia and in Thailand. Our findings show that the wage returns to the investment in job training decrease significantly as one controls for worker's and firm's characteristics. We find that on-the-job training is associated with increases in individual wages of 7.7% in Malaysia and 4.5% in Thailand. We also estimate that wage returns to on-the-job training tend to be quantitatively higher for men than for women although in Thailand they are not statistically significant for males. In Malaysia, the returns for males are 11% while for women they are not statistically

different from zero. Workers that have at least completed secondary education also report higher returns to on-the-job training than other workers (returns are 9% and 10% for Malaysia and Thailand, respectively).

In the empirical work, we start from a simple worker level Mincer type equation relating hourly wages with several observable worker and firm characteristics, including differences in the incidence of on-the-job training. Our main coefficient of interest quantifies the average effect on wages of having received on-the-job training. However, the estimation of the effect of on-the-job training on wages poses a major challenge as training is likely to be an endogenous variable to wages. On-the-job training is a choice variable for both firms and workers and most likely is also correlated with worker and firm characteristics, which in turn are also correlated with labor productivity and wages. Failure to control in a flexible manner for these characteristics may create a bias in the estimates of the effect of training on wages, as workers selecting into training may have different characteristics. In our empirical approach, we hope to minimize this problem by exploring a rich data set with many worker and firm characteristics and the propensity score matching (PSM) method. When compared to ordinary least squares (OLS), the PSM estimates allow for a more flexible (non-linear) functional form relating observable worker characteristics and their wages.

The propensity to score matching method is developed in two steps. First, it estimates the probability of each worker to be selected into the training provided by the firm, given his or her observable characteristics. Based on this probability, it generates a “control group” of workers that did not participate in the training but whose probability of being selected into on-the-job training was very similar to the probability observed for the sample of trained individuals. These workers are very “similar” to those actually participating in training in all their observable characteristics (e.g., education, occupation in the labor market, years of experience). The only thing that distinguishes them from the trained workers is the sole fact of not having received on-the-job training. Hence, the wage difference between these two groups can be fully attributed to the wage impacts of on-the-job training.

Our empirical findings document two interesting patterns across the two countries. First, the incidence of on-the-job training differs significantly by several worker and firm characteristics. In particular, we find that the more educated and more tenured workers are, in both countries, the more likely they are to receive on-the-job training. We also show that larger, more innovative foreign firms are also more likely to invest in on-the-job training. Second, there is strong evidence that the workers’ wages increase with the incidence of on-the-job training in both countries. In our preferred estimates exploring propensity score matching, the average wage returns to on-the-job training are 7.7% in Malaysia and 4.5% in Thailand. Furthermore, the heterogeneity analysis shows that in Malaysia the wage returns to job training are larger for men (11%) than for women (for whom they are not statistically different from zero). We also find that in both countries there are higher wage returns to job training for workers with completed secondary education or more years of education when compared to those who have not completed secondary education. The returns to on-the-job training for workers with at least secondary education are 9% in Malaysia and 10% in Thailand, respectively. In contrast, for workers with lower levels of education, there is no evidence of positive wage impacts both for Malaysia and Thailand. These findings clearly reinforce the idea that the investment in job training is complementary to the initial level of education of workers.

Our paper relates closely to two empirical literatures. First, it relates to the work analyzing the firm's investment in on-the-job training in developing countries (e.g., Ariga and Brunello, 2002, Almeida and Aterido, 2010, 2011, Almeida, 2010 and Almeida and Cho, 2012). The main empirical patterns found in these papers for developing countries are close to the findings found for developed countries (e.g., Bassinini et al. 2005). Larger, more open and innovative firms, with a more skilled workforce and operating in more technologically intensive sectors are more likely to train their employees. The major exception is Ariga and Brunello (2002). Exploring an employee survey for Thailand in 2001, they find a significant and negative relationship between years of formal educational and training.

Second, we relate to the empirical work quantifying the wage returns to on-the-job training exploring worker level data⁵. Table A1 in the Appendix summarizes some of the main empirical studies quantifying the wage returns to on-the-job training, for developing and developed countries. Panels A and B report the estimates from papers using worker level data. Panel A refers to developed countries and Panel B to developing countries⁶. A word of caution is needed when comparing cross country estimates of the returns to on-the-job training. First, the variable capturing on-the-job training differs significantly across data sets yielding reduced comparability across studies. Second, there is little comparability in the reduced form equation used across most of the analysis⁷.

The point estimates reported in Panel A for developed countries are very diverse. Some studies report positive and significant wage returns to training. However, more recently, as longitudinal data becomes available and experimental methods are used, the wage returns to on-the-job training tend to be smaller than in the cross section studies. Furthermore, in some cases, the returns are even zero (e.g., Leuven and Oosterbeek, 2002, 2004)⁸. The point estimates for most developing countries, reported in Panel B, are generally in the order of 20%. The evidence in the panel is also quite diverse. Chung (2000) and Johanson and Wanga (2008) explore cross sectional data and find evidence of large returns (between 20% and 38%) for Malaysia and Tanzania, respectively. On the other hand, Frazer (2006) finds that in Ghana, during the 90s, the returns to apprenticeship training were not statistically different from zero. Monk et al. (2008) find in addition some heterogeneity within the country and across education levels. They show that the returns of apprenticeships are 50% for individuals with no education but decline as education raises. They find evidence that the returns are zero for individuals with more than 6 years of formal education.

The methodology in our paper is closer to Rosholm et al. (2007). They estimate that the returns to training are on average 21% for Kenya and that in Zambia training is not associated with higher wages. Like us, they also explore a matched employer and employee data set (collected by the World Bank) and a propensity score matching methodology. However, the larger number of observations in our sample and the more detailed information on worker and firm characteristics allowed us to conduct a deeper analysis. First, we consider hourly wages as a dependent variable, while Rosholm et al. (2005) consider only monthly wages. Second, we are able to control both for detailed worker human capital characteristics and for several firm characteristics that they do not. At the worker level, we include variables such as having received training at a previous employer, owning a bank account and using the internet. These variables,

especially past training, will prove to be important in explaining the selection into training. At the firm level, we are able to control for the average years of schooling of the workforce, for the degree of innovation or for the degree of exports.

The paper proceeds as follows. Section 2 describes the dataset used and the descriptive statistics. Particularly, in Section 2.1 we explain in detail our main dependent variable of interest: the logarithm of workers' hourly wages. In Section 3 we analyze which variables determine the selection into training. Section 4 presents the propensity score matching estimates for the wage returns to on-the-job training. In section 4.1, we explain the empirical model, and in section 4.2 we report the main empirical findings for the wage returns to on-the-job training. In Section 5, we report heterogeneity analysis by gender and level of education. Finally, Section 6 concludes.

2. Data and descriptive statistics

We explore a matched employer-employee data set collected by the World Bank, *Enterprise Surveys*, for Malaysia (2002) and Thailand (2004)⁹. A total of 1,152 firms were surveyed in Malaysia and 1,385 in Thailand. For each firm, a random sample of 10 employees in each firm was interviewed yielding a total of 10,822 and 13,850 firm-worker observations in Malaysia and Thailand, respectively. However, in the analysis we have excluded observations with missing values for the main covariates of interest both at the firm and worker level. As a result, the number of observations used will be 6,679 for Malaysia and to 9,418 for Thailand, respectively.

This data set has several advantages to study this topic. First, the questionnaire is similar across the two countries, which ensures comparability of the results. Second, the survey collects simultaneously detailed information on worker and firm characteristics. In particular, at the firm level it collects information on the sector of activity, geographical location¹⁰, total number of employees, public and foreign ownership as well as information on the human capital of the manager, on the average years of formal education of the workforce, number of employees per occupation, and percentage of women in the firm. The survey also gathers information on technological variables or investments in new production technologies such as R&D expenses, introduction of new products and adoption of new technologies. At the worker level, it collects information on gender, age, marital status and nationality. Most importantly it also collects detailed human capital characteristics like years of formal education, tenure with the firm, years of experience in the labor market, and whether each worker enrolled in vocational training programs in the past. Finally, the survey collects information on whether the firm offered on-the-job training to their employees last year and whether the employees interviewed took any formal training since they joined that firm. In addition, monthly wages and hours of work per week are also reported.

In particular, the survey contains the following information about formal training programs at the firm and at the worker level. At the firm level, the survey asks: "*Did your plant run formal in-house training programs for its employees in 2001?*", "*Did your plant send employees to formal training programs run by other organizations during the fiscal year of 2001?*" At the worker level the survey asks: "*Have you received formal training since you joined this firm?*". Based on these two questions, we constructed two variables capturing the incidence of on-the-job training at the firm and at the worker level. First, we constructed a firm level dummy variable that equals one if the firm

offered formal training to its workers in the year prior to the survey. Second, we constructed a worker level dummy variable that equals one if the worker has received formal training since he joined that firm. In addition, for those workers whose the current position is not their first job, we have information on whether the worker received training at his previous job¹¹.

Table A2 in the Appendix describes the main variables used. Tables A3 and A4 in the Appendix report summary statistics for the main firm and worker characteristics used in the paper. In Malaysia, the final sample covers manufacturing (79%) and services (21%). In Thailand, the sample only covers manufacturing. In addition, the distribution of firms across the two countries is different. While in Malaysia, small firms are approximately half of the sample, in Thailand, medium, large and very large firms account for more than 70% of the sample. In the two countries, approximately 70% of the firms are domestic owned and a large share exports at least some of their sales (66% in Thailand and 62% in Malaysia). Rubber and Plastics (22%) and food processing (18%) are the two more represented industries in Malaysia. In the Thai sample, firms are more equally divided among the different sectors. Finally, firms in Malaysia have a higher share of skilled labor (49%) than in Thailand (24%) and the average of years of formal education is also slightly higher in Malaysia than in Thailand.

Finally, Table A3 also shows that the training incidence at the firm level is 51% in Malaysia and 76% in Thailand¹². The incidence of training is smaller in Malaysia in part due to the low training incidence of job training among the firms operating in Rubber and Plastics (36% of the firms train) and in Food Processing (55% of the firms train). Also interestingly, in Malaysia most of the firms that offer training explore both facilities in house and externally. Most of the training costs are supported directly by the firms (at least formally, as firms can transfer the cost of training to employees through lower wages). Only 6% of the firms for Malaysia and 3% for Thailand report to have shared the costs of training with their employees¹³.

Table A4 in the Appendix computes summary statistics for the sample of workers in both countries. In both samples women represent approximately half of the sample. The average age, tenure and years of experience is also quite similar across the two countries. Again, the human capital of the workforce is higher in Malaysia than in Thailand. In Malaysia only 15% of the workers have up to primary education compared with 30% of the workers in Thailand. In Malaysia there are also more workers with polytechnic or vocational education than in Thailand (15% vs. 6%). This higher human capital translates also into more skilled occupations in Malaysia than in Thailand in our sample. While skilled production workers is the most represented group in the Malaysian sample (36%), in Thailand, the most represented occupation group is unskilled production workers (37%). Also interestingly, Malaysian workers have been more exposed to foreign languages and cultures. In particular, 7% of the Malaysian workers but less than 1% of Thai workers studied in a foreign country.

Table A4 in the Appendix also shows that the incidence of on-the-job training is higher in Thailand than in Malaysia also at the worker level. In Malaysia, 33% of the workers report having received some training since they joined the firm. In Thailand, this number is 52% of the workforce. The percentage of workers that received training at the previous employer was 17% in Malaysia and 24% in Thailand.

2.1. Dependent variable

The main dependent variable of interest is the natural logarithm of the hourly wage (in USD)¹⁴. In Table 1 we present the average log hourly wage as well as the raw difference in the average hourly wage for both trained and not trained workers.

The results in the table show that workers who report having received some formal on-the-job training since joining the firm report higher earnings than non-trainees in both countries. However, this difference in the average wages is likely capturing the effect of worker and firm characteristics that drive the selection of workers into training and that, simultaneously, also influence their hourly wages.

3. The selection of firms and workers into training

In order to understand which variables influence the selection into on-the-job training we run a series of regressions at the firm and worker level. We assume that firms decide whether or not to train their workers if the profits from this investment are greater than the costs:

$$Train_{jfr} = \begin{cases} 1 & \text{if } \pi^*_{jfr} > 0 \\ 0 & \text{otherwise.} \end{cases} \quad (1)$$

where $Train_{jfr}$ is a dummy variable that equals one if firm j , operating in industry f and region r offered on-the-job training to its employees during the year prior to the survey and π^*_{jfr} are the net benefits of investing in training. Since π^*_{jfr} is unobservable we

Table 1 Differences in wages for workers with and without training

	Mean (1)	SE (2)
Panel A: Malaysia		
Avg Ln wage	0.410	(0.778)
Avg Ln wage for the trained	0.698	(0.793)
Avg Ln wage for the non-trained	0.266	(0.729)
Difference in Avg Ln wage for the trained and the not trained	0.431***	(0.019)
Number of workers trained		2,215
Number of workers not trained		4,464
Total number of workers		6,679
Panel B: Thailand		
Avg Ln wage	-0.135	(0.715)
Avg Ln wage for the trained	0.000	(0.743)
Avg Ln wage for the non-trained	-0.285	(0.657)
Difference in Avg Ln wage for the trained and the not trained	0.284***	(0.014)
Number of workers trained		4,941
Number of workers not trained		4,477
Total number of workers		9,418

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: We report the natural logarithm of hourly wage in USD at the time of the survey. * significant at 10%, ** significant at 5%, *** significant at 1%. Panel A reports the descriptive statistics for the sample of workers in Malaysia and Panel B reports the descriptive statistics for the sample of workers in Thailand. Column (1) reports the mean of the variables listed, and column (2) and reports standard errors.

assume π_{jfr}^* is a function of several firm, industry and regional characteristics. We also assume that π_{jfr}^* is linear so that $\pi_{jfr}^* = \delta Z_{jfr} + \mu_f + \mu_r + \varepsilon_{jfr}$, where Z_{jfr} is a vector of firm characteristics and μ_f are industry fixed effects, μ_r are region fixed effects, and ε_{jfr} captures unobserved firm characteristics. Given this linear form, the probability that firm j offers formal on-the-job training to its employees is given by:

$$\Pr(\text{Train}_{jfr} = 1) = \Pr(v_{jfr} > -\delta Z_{jfr} - \mu_f - \mu_r) \quad (2)$$

Assuming that the error term ε_{jfr} follows a normal distribution, equation (2) can be estimated by maximum likelihood (probit). Tables A5 and A6 in the Appendix report the estimates of different specifications of equation (2) in the text for Malaysia and Thailand, respectively. Specifications (1) through (6) differ in the observable firm characteristics that are included. In all specifications we control for two-digit ISIC industry codes and for region dummies. Specification (7), which we will consider our baseline specification of firm characteristics, includes the interaction of industry and region fixed effects. In this specification, we control for size of the firm, age, export intensity, foreign ownership, education of the workforce (including managerial education) and degree of technological adoption. The findings show that training incidence increases with firm size in both countries although not with age of the firm. Training incidence increases also with the firm's presence in external markets and with foreign ownership. For example, in Malaysia, training incidence is 56.4 percentage points higher in a firm with more than 250 employees than in a micro firm. We also find robust evidence that training incidence increases with the human capital of the workforce (measured both with years of education, skills of the workforce and by managerial education) and with the degree of technological adoption in the firm.

In Tables A7 and A8 we replicate the estimation of equation (2) with maximum likelihood (probit model) but consider on-the-job training incidence at the worker level as the dependent variable¹⁵. This analysis is a critical first step of the propensity score matching methodology. We consider several observable characteristics that likely determine the selection into training and that we can quantify with our detailed data set. We will then estimate the fitted values for each worker level observation. Therefore, for each worker who has received training it is feasible to match him/her with a worker with a close enough fitted probability. This group of workers will constitute the control group in the estimation of the impacts of on-the-job training on wages.

Specifications (1) through (5) of Tables A7 and A8 always include the baseline firm characteristics reported in column (7) of Tables A5 and A6. However, the set of worker level characteristics differs across columns. Column (1), in addition to the baseline firm characteristics, controls for the worker's education (including vocational education), gender, age, tenure with the firm, potential experience, marital status, occupation, if the worker is an apprentice and if he belongs to a trade union. In columns (2) through (5) we add dummy variables capturing if the worker has a computer at home (specification 2), owns a bank account (specification 3), uses regularly internet for transactions (specification 4) and has received training at a previous employer (specification 5). In Table 2, we report the results for both countries exploring our preferred specification (specification (5) in Tables A7 and A8 in the Appendix).

Table 2 Average marginal effects and asymptotic errors from the worker level probit for participation in training

	Malaysia		Thailand	
	Coefficient (1)	SE (2)	Coefficient (3)	SE (4)
Degree	0.523***	[0.186]	0.426**	[0.215]
Diploma	0.608***	[0.180]	0.353	[0.215]
Upper secondary	0.458***	[0.172]	0.226	[0.212]
Lower secondary	0.266	[0.172]	0.114	[0.211]
Primary	0.036	[0.176]	-0.034	[0.209]
Vocational education	0.101**	[0.049]	-0.015	[0.059]
Woman	-0.071	[0.038]	-0.072*	[0.030]
Age of worker	0.020	[0.019]	0.020	[0.016]
Tenure with the firm	0.069***	[0.001]	0.100***	[0.008]
Potential experience	-0.032***	[0.009]	-0.003	[0.008]
Single	-0.146***	[0.045]	0.022	[0.033]
Manager	0.060	[0.066]	-0.087	[0.070]
Professional	0.118	[0.078]	-0.055	[0.059]
Skill production	-0.017	[0.054]	0.181***	[0.046]
Unskilled production	-0.269***	[0.062]	0.225***	[0.047]
Apprentice	0.506**	[0.204]	-0.234	[0.174]
Unionized	0.531***	[0.088]	0.628***	[0.144]
Computer	0.108***	[0.040]	0.143***	[0.040]
Bank account	-0.040	[0.095]	0.475***	[0.052]
Transaction internet	0.277***	[0.095]	0.193	[0.141]
Training at the previous firm	0.793***	[0.048]	0.607***	[0.035]
Firm base specification?	Yes		Yes	
Observations	6,679		9,418	

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: The dependent variable is a dummy variable that assumes the value 1 if the worker received any formal training after joining the firm. The table reports the marginal effects (at mean values) on the firm's propensity to train from probit regressions (equation 3 in the text). The regressions control for several firm level characteristics, as it is listed in the specification reported in column (7) of Tables A5 and A6 of the Appendix. * significant at 10%, ** significant at 5%, *** significant at 1%. Columns (1) and (3) report the coefficient of the variable, and columns (2) and (4) report standard errors. These are the same coefficients as column (5) in Tables A7 and A8 of the Appendix.

The findings show that training incidence increases with the level of human capital of the worker from secondary education onwards and, interestingly, also increases as individuals hold some degree of vocational training. Women are less likely to receive on-the-job training in Thailand but not in Malaysia. Workers with longer tenure with the firm are more likely to receive on-the-job training in both countries. Differences in the probability to participate in training are not statistically significant when comparing workers of different ages. In Thailand, training incidence is higher both for skilled and unskilled production workers than for non-production workers. In Malaysia, we only find that unskilled production workers are less likely to train than non-production workers. This might be a result driven by the industries represented in the sample¹⁶. In both countries, we find that workers that belong to a trade union and use a computer

at home are more likely to participate in job training than others. In Malaysia, workers that have ever made a transaction through the internet also tend to be selected into training more frequently than those who have never used e-commerce. Moreover, we also find that current training incidence is very strongly and positively correlated with past training incidence for both countries.

4. The wage returns to on-the-job training

The accumulation of human capital has long been seen as an investment decision (Becker, 1964). While investing, each individual gives up some proportion of income during the education and training period in exchange of increased future earnings. Individuals will be willing to take additional schooling or training if the costs (tuition and training course fees, forgone earnings while at school and reduced wages during the training period) are compensated by higher future earnings. Assuming perfectly competitive labor markets, wages reflect the marginal product of workers and should increase with the accumulation of human capital if individuals become more productive in their current job¹⁷.

4.1. Propensity score matching

We use propensity score matching to quantify the wage returns of job training at the worker level. The main idea underneath the propensity to score matching methodology is to match as closely as possible individuals who have received training to those not receiving on-the-job training, to then be able to meaningfully compare the differences in their wages¹⁸.

Let $W_{1i} - W_{0i}$ (4) be the difference between a worker's log wage W_i that results from participating in training. If worker i participates in training, he gets wage W_{1i} and if he does not receive training, his wage is W_{0i} ($D_i \in \{0, 1\}$ summarizes worker i 's treatment, 1 and 0 mean, respectively, receiving and not receiving training). We do not observe, however, the wages of individuals who receive training if they had not received it, that is W_{0i} , and we do not observe wages for the individuals who were not trained if they were trained W_{1i} .

As we have argued in a previous section, to assess the impact of job training it is not enough to compute the average wage difference between the workers that received training and the ones that did not participate in training. The main reason is that most likely a large part of this difference can be caused not by training itself but also be driven by other worker and firm characteristics that determine the selection into training. The propensity to score matching (PSM) methodology departs from the assumption that all the relevant differences between the treated and the untreated individuals are captured by their observables. Within the group of the untreated it selects a group as similar as possible to the treated group. The difference in wages across workers who received training and this set of workers is a better estimate of the returns to training.

First, given the richness of our data set, we assume that a significant number of worker and firm variables (X) explain the relevant differences between the treated and the untreated groups. For consistent estimations it is required that $W_{0i} \perp D_i | X$ where X is the set of observed variables¹⁹. However, if X is multidimensional it becomes difficult to match the individuals. Rosenbaum and Rubin (1983) have proved that $W_{0i} \perp D_i | X$ implies $W_{0i} \perp D_i | p(X)$, where $p(X)$ are the propensity score fitted values, or the probability of participating in

training. Therefore the untreated individuals that present higher probabilities of receiving training will compose the counterfactual group.

The first step of this method is thus to estimate the probability of each worker to receive on-the-job training. This is given by the fitted values of the worker level probit regressions for the incidence of training. We consider the specification reported in Table 2. There, we control for several worker (including education, gender, age, tenure with the firm, potential experience, marital status, occupation) and firm characteristics (including firm size, age, export intensity, foreign ownership, education of the workforce, degree of technological adoption). The reason why we also include firm level characteristics is to control for the fact that the training decision depends partly on the firm. By matching workers with similar characteristics and who work for similar firms we hope to minimize the selection bias that is likely arising from the fact that individuals selected into training may be the ones with higher unobserved ability.

Because the treated individual i can be matched with one or n individuals on the non-treated group, we choose the one-to- n matching method. This implies that each individual in the “treatment” group is matched with a weighted average of all individuals in the “control” group that have similar fitted values²⁰. After associating each treated individual i with a mean of untreated individuals with different weights we simply compute the difference between the averages of the log wages in the treated group and in the control “weighted average” to quantify the causal effect of on-the-job training on wages.

In Malaysia there is a set of 4,425 untreated individuals and 2,202 treated (a total support group of 6,627). In Thailand, the set of untreated and treated groups have 4,477 and 4,941 individuals, respectively (9,418 individuals in total). We compute the average treatment effect on the treated individuals (ATT), yielding the impact of the training on the set of workers who actually end up receiving it. Table A11, in the Appendix, reports the balancing tests to check the quality of the matching methodology to our sample. It is reassuring to see statistically similar means in most of the covariates for both the treated and the control groups. The only exception is the variable tenure in Malaysia for which the t-statistic rejects the null hypothesis.

4.2. Empirical results for the returns to training

As mentioned above workers selected into training may receive on average different wages than those not selected. In this section, we present estimates of the average treatment effect on the treated (ATT), which is the effect of the training (received since joining the firm) on the hourly log wages of trained workers. We report both the raw log wage differences and the ATT using propensity score matching. Table 3 reports that the raw difference in wages between the treated and the untreated groups is 42.9% for Malaysia and 28.4% for Thailand. Once we explore the propensity to score matching methodology, we find that the impact of on-the-job training on hourly wages falls to 7.7% for Malaysia and 4.5% for Thailand, respectively. The estimates are significant at a 5% level of confidence for both countries.

Therefore, Malaysia presents higher returns from job training than Thailand. A priori, this is not immediate. On the one hand, Malaysia has a higher per capita gross domestic product and also has more youth in schools than Thailand, suggesting that their returns to human capital could be smaller than in Thailand. On the other hand, the accumulated stock of capital in Malaysia is higher, and if skills and capital are

Table 3 Wage returns to on-the-job training

	Difference in log wages (1)	SE (2)	t-stat (3)
Panel A: Malaysia			
Unmatched	0.429***	0.02	21.9
ATT	0.077***	0.03	2.92
Untreated	4,425		
Treated	2,202		
Sample size	6,627		
Panel B: Thailand			
Unmatched	0.284***	0.02	15.46
ATT	0.045**	0.02	2.37
Untreated	4,477		
Treated	4,941		
Sample size	9,418		

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: * significant at 10%, ** significant at 5%, *** significant at 1%. The table explores the propensity score matching estimate of equation (4) in the text. Panel A reports the PSM results for the sample of workers in Malaysia, and Panel B reports the results for the sample of workers in Thailand. Column (1) reports the Average Treatment Effect on the Treated (ATT). The estimate gives the impact on wages of training for those actually participating in training. Column (2) reports standard errors and column (3) reports the t-statistic.

complementary, all else constant, the returns to human capital could be higher than in Thailand. Also if training presents decreasing returns, it is reassuring to see that returns are lower in the country with the highest training incidence: Thailand.

For comparability, in Tables A9 and A10 in the Appendix, we show the estimates for the impact of training on wages, using the least squares methodology (OLS) and considering alternative specifications. The specification that is closer to the variables we control in the PSM estimates is reported in column (3). Comparing columns (1) and (3) in Tables A9 and A10, we see that the wage difference between trainees and non-trainees falls from 43.1% to 4.3% in Malaysia and from 28.4% to 4.2% in Thailand. Even though the numbers are very similar for Thailand and for Malaysia, the OLS estimates are lower than the PSM estimates, suggesting that least squares estimates have a downward bias.

5. Heterogeneity analysis

Until now we assumed that returns to on-the-job training are the same for all the workers and firms within each country. In this section, we allow for the returns to be different by two fundamental worker characteristics: gender and education. Our results show that the returns to training are higher for men than for women in Malaysia, and for workers with completed secondary education or more education in both countries, when compared with workers that have not completed secondary education.

Table 4 reports the main results for the sample of men and women, separately. Panel A reports the results for Malaysia, while Panel B reports the results for Thailand. Table 4 shows that, in Malaysia, wage returns are higher for men than for women. Men present wage returns to on-the-job training of 11%, while there are no statistically significant

Table 4 Wage returns to on-the-job training, by gender

	Sample of men			Sample of women		
	Difference in log wages	SE	t-stat	Difference in log wages	SE	t-stat
	(1)	(2)	(3)	(4)	(5)	(6)
Panel A: Malaysia						
ATT	0.11***	0.04	2.95	0.05	0.04	1.30
Untreated	2,394			2,035		
Treated	1,269			933		
Observations	3,663			2,968		
Panel B: Thailand						
ATT	0.05	0.03	1.44	0.04*	0.02	1.81
Untreated	1,970			2,507		
Treated	2,182			2,759		
Observations	3,663			5,266		

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: * significant at 10%, ** significant at 5%, *** significant at 1%. The table uses propensity score matching to estimate equation (4) in the text. We estimate separate regressions by gender. Columns (1) and (4) report Average Treatment Effect on the Treated (ATT) which evaluates the wage impact of training for those actually participating in training. Columns (2) and (5) report standard errors. Columns (3) and (6) report the t-statistic. Treated individuals are those who have participated in training and the untreated individuals are the "control group" that is similar for all characteristics to the treated group except for the fact of receiving training. Panel A reports the estimates for the sample of workers in Malaysia, and Panel B reports the estimates for the sample of workers in Thailand.

returns for women. This may be explained by the fact that women tend to go into and out of the job market more frequently than men and thus may be less likely to receive on-the-job training. This higher turnover may also make it difficult to appropriate the returns from the investments in job training. We do not find, however, that same result for Thailand. There the wage returns to job training are quantitatively larger for men than for women, although for men they are not statistically different from zero (at a 10% level).

Table 5 reports the estimates for the samples of workers that have completed at least secondary education and for the sample of workers with less education. Panel A shows the results for Malaysia, and Panel B shows the results for Thailand. Consistent also with the findings of others in the literature, we show that the returns to on-the-job training are higher for workers that have completed at least secondary education than for workers with less education. Workers with completed secondary education have returns to training of 9% in Malaysia and of 10% in Thailand. In contrast, for those workers without completed secondary schooling, there are no statistically significant returns to on-the-job training in both countries.

6. Conclusion

In developing countries, governments are increasingly concerned with the rapidly changing demand for skills and the slow response of the general and vocational schooling tracks. As a consequence, many employers complain about the lack of skills and education of their workforce. Policymakers are thus increasingly concerned that the supply of skills in the market does not keep pace with the demand and think about the design of policies to address this problem. The investment in on-the-job training is one important way to mitigate this gap by developing job relevant skills among the workforce.

Table 5 Estimated wage returns to job training, by level of education

	Workers completed secondary education or more years of schooling			Workers with up to incomplete secondary education		
	Difference in log wages	SE	t-stat	Difference in log wages	SE	t-stat
	(1)	(2)	(3)	(4)	(5)	(6)
Panel A: Malaysia						
ATT	0.09***	0.03	2.61	0.05	0.04	1.11
Untreated	2,075			2,242		
Treated	1,642			548		
Observations	3,717			2,790		
Panel B: Thailand						
ATT	0.10**	0.04	2.51	0.03	0.02	1.50
Untreated	1,016			3,431		
Treated	1,935			2,992		
Observations	2,951			6,423		

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: * significant at 10%, ** significant at 5%, *** significant at 1%. The table uses propensity score matching to estimate equation (4) in the text. We estimate separate regressions by education group. Columns (1) to (3) refer to the sample of workers that have completed secondary education or more. Columns (4) through (6) refer to workers with lower levels of education (that is those with up to incomplete secondary education). Columns (1) and (4) report ATT (Average Treatment Effect on the Treated), it evaluates the wage impact of training for those actually participating in training. Columns (2) and (5) report standard errors, and columns (3) and (6) report the t-statistic. Treated individuals are those who have participated in training, and the untreated individuals are the "control group" that is similar for all characteristics to the treated group except for the fact of receiving training. Panel A reports the estimates for the sample of workers in Malaysia, and Panel B reports the estimates for the sample of workers in Thailand.

The evidence on this topic is generally scarce for developing countries. The measurement of returns to training presents several challenges, and this is the reason why we find so different results in the literature. Variables are usually not comparable across studies, and sometimes data-sets do not allow for an accurate estimation of the results.

In this paper we quantify the wage returns from on-the-job training in Malaysia and in Thailand exploring a unique data set matching workers and firms. Using a matching estimators method to control for the selection bias we find returns of 7.7% and 4.5% for Malaysia and Thailand, respectively. In Malaysia, we find that returns are clearly higher for men than for women. Workers that have completed secondary education or more also show higher wage returns than those who have not completed secondary schooling. Economic theory tells us that the wage effects are a lower bound estimate for the effect of training in productivity. Therefore the productivity impact of training in these countries should be even higher than the estimated values.

Endnotes

¹We will explore the *Enterprise Surveys* collected by the World Bank. In each country the survey inquires whether the firms invested in formal on-the-job training programs. This paper restricts the attention to employer provided formal on-the-job training programs. It will not address informal job relevant training. Johanson and Wanga (2008) present evidence for wage returns to training in the informal sector. Kahyarara and Teal (2008) discuss the link between other types of training and labor

market outcomes by examining the wage returns to vocational training. Fitzenberger et al. Völter (2007) study the effects of public training in helping the transitions between unemployment and employment, and Frazer (2005) studies the effectiveness of training for the self-employed in developing countries.

²These figures mean that Thailand presents a per capita GDP similar to countries such as Peru, El Salvador and Tunisia whereas Malaysia is closer to countries like Latvia Lithuania and Chile.

³See Panel B of Table A1 in the Appendix for a summary of the papers that have analyzed wage returns to training in developing countries.

⁴There are some empirical papers that use firm level data to try evaluate the impact of offered training on the firms' productivity (e.g. Almeida and Carneiro 2008; Barret and O'Connell 1999) and on average wages (Lopez-Acevedo and Tan 2003).

⁵Few empirical papers have looked at the extent to which the benefits of training (ultimately effects on higher firm productivity) are shared with workers. One exception is Dearden et al. (2006) for the UK.

⁶Panels C and D summarize the works using firm and industry level data.

⁷For example, some papers have defined training incidence with a dummy variable capturing whether training was offered over the previous year to the survey. Others, like Bassinini et al. (2005) use the accumulated stock of training hours over the sample period (6 years). Moreover, the reduced form estimated typically depends on the data available, which in turn differs across data sets and countries. For a similar point see Haelermans and Borghans (2012).

⁸Leuven and Oosterbeek (2002, 2004) use two different methods to estimate the returns to training in Holland. Leuven and Oosterbeek (2002) identify individuals planning to enroll in a training program but did not do so due to a random event and find evidence of no returns to job training. Leuven and Oosterbeek (2004) explore a discontinuity which allowed firms to reduce their training expenses only for workers more than 40 years old Although their results are just valid locally, they also conclude that there were no returns from the investment in job training. Similarly, Sousounis (2009) explores longitudinal data and does not find evidence that training increases wages in the U.K. between 1998 and 2005.

⁹The information collected in the *Enterprise Surveys* is based on one to two hour interviews with the firm manager. This data set has been used for studying this and other topics (see e.g. Almeida and Aterido, 2011, Almeida and Carneiro, 2008a, Almeida and Fernandes, 2008). Previous versions of this project within the World Bank include the *Regional Program on Enterprise Development* collecting firm and worker level data in Sub-Saharan Africa countries for a decade (e.g., Rosholm et al. 2007, Frazer 2006) and the *World Business Environment Survey*.

¹⁰In the Malaysian sample we have firms from the Central Region: Selangor, KL, Melaka (4,641 observations); the North Region: Penang, Kedah (1,899 observations); the South Region: Johor (3,290 observations); the East Coast: Terengganu (181 observations); the Northeast (320 observations) and the South (390 observations). In Thailand firms in the sample operate in the North (730 observations); the Centre (3,260 observations); Bangkok and Vicinity (6,160 observations); the East (1,920 observations); the Northeast (320 observations) and the South (390 observations).

¹¹The survey asks: "Did you receive formal training at the previous employer?"

¹²This difference is not driven by the manufacturing/services balance in our sample. The training incidence in Malaysia is 49% only for the manufacturing sector.

¹³In particular, firms are asked “Did the employees share the cost of training?”

¹⁴We report log hourly wages in USD in 2002 prices for Malaysia and 2004 prices for Thailand.

¹⁵We assume that a firm offers formal training to a worker if there is a net positive benefit of this investment. The main difference is that now the benefits should also be a function of the worker level observable characteristics, captured by X_{ijfr} . In this case, the probability of a worker i being employed in firm j is determined by his characteristics (X_{ijfr}) and the firm characteristics (Z_{jfr}) so that $\Pr(\text{Train}_{ijfr} = 1) = \Pr(v_{ijfr} > X_{ijfr} - \delta Z_{jfr} - \mu_f - \mu_r)$ (3).

¹⁶As explained in section 2 of the text our sample only includes formal manufacturing firms.

¹⁷With imperfect competition wages do not necessarily reflect labor productivity and therefore might not reflect changes in the worker’s productivity.

¹⁸For another application of this method see Rosholm et al. (2007), and for a complete theoretical discussion of the matching estimators see Heckman et al. (1999).

¹⁹This is defined, in the econometric literature, as Conditional Independence Assumption (CIA).

²⁰A normal kernel is used to define the weights.

²¹The worker variables included were: educational attainment, gender, age, tenure in the firm, years of labor market experience, marital status, occupation, whether the individual is member of a labor union, owns a computer, a bank account, has ever made an internet transaction and whether the worker received training at a previous employer.

²²In addition to the worker variables we described in the previous footnote we include the following firm characteristics: size, foreign capital participation, exports, average years of education of the work force, education of the manager, introduction of new production technologies, industry and region.

²³Tables A9 and A10, in the Appendix, report the results for the worker variables included in the regressions. We focus on the findings in column (3). The estimates show that, in both countries, the returns to schooling are increasing with the level of formal education completed. Women and unionized workers earn lower wages than men and non-unionized workers. Wages also tend to increase with age, tenure and experience. Moreover, wages for managers and professionals are higher than the wages of non-production workers (omitted occupation group) and skilled production, unskilled production in both countries. Finally, those that report having a computer at home, a bank account, and using the internet regularly also report higher wages. The same happens for those individuals reporting having received training with their previous employer. Perhaps surprisingly, in Malaysia, we find that returns from past training are higher than returns from more recent training events (6.8% vs. 4.3%), as we expect training to depreciate with time.

Appendix

A1. Additional tables (Tables A1, A2, A3, A4, A5, A6, A7, A8, A9, A10, A11)

Table A1 Review of the literature on wage returns of job training

Panel A: Papers using worker level data - developed countries

Name of study	Data, country and time period	Dependent variable	Training	Other independent variables	Effects training on wages	Controls endogeneity training
Bassinini, Booth, Brunello, De Paola and Leuven (2005)	Worker Level Data for Europe (1995-2001)	log hourly wage	sum over the sample period (6 years) of training events	age, gender, marital status, schooling, year, country and industry.	Between 3.7% and 21.6% depending on the country.	No
Albert, Serrano, Hernanz (2010)	Worker Level Data for Europe (1995-2001)	log hourly wage	dummy: having participated in training between January of the previous year and the date of the interview	gender, educational attainment, potential labour market experience, firm size, industry affiliation, working time, occupation, seniority	Positive for OLS but not statistically significant for FE.	Yes Fixed effects
Budría and Pereira (2004)	Worker Level Data for Portugal (1998-2000)	log hourly wage	dummy: having ever participated in training	age, experience, schooling (and interactions between these variables and training), part time, tenure, sector, firm size.	12% for men and 37% for women	Yes Excluded instruments (selection model): having a second activity and having lived abroad.
Dearden, Reed and Reenen (2006)	Worker Level Data for U.K. (1992,1997)	log hourly wage	dummies: having participated in training in the previous 4, 3, 2, and 1 quarters	age, gender, occupation, dummy for no qualifications, firm size, industry.	0.15% (production sector only)	No
Leuven and Oosterbeek (2002)	Worker Level Data for Holland (2001)	log hourly wage	dummy: having participated in training in the previous 12 months	age, gender, schooling, firm size.	Not statistically different from zero	Yes Randomization: control group composed by people that were planning to engage in a training activity by did not because of some random event.
Leuven and Oosterbeek (2004)	Worker Level Data for Holland (1999)	log hourly wage	dummy: having participated in training in the previous 12 months	age, schooling, tenure, firm size.	Not statistically different from zero (for 40-year-old workers)	Yes They use the RD data design method. They explore the discontinuity introduced by a new tax law that allows tax deduction for firms' expenditures on training for workers with more than 40 years. So the decision of training workers around age 40 suffers and will be influenced by an exogenous effect (the law).

Table A1 Review of the literature on wage returns of job training (Continued)

Lillard and Tan (1986)	Worker Level Data for U.S.A. (1983)	log annual wage	dummies: having participated in training (formal and informal) in the current job	experience, schooling, tenure, union member, dummy for non-white, tenure, region, long run state unemployment rate, cyclical sensitivity of state unemployment.	22% for formal training	No
Sousounis (2009)	Worker Level Data for U.K. (1998-2005)	log weekly wage	dummy: having participated in training in the previous 12 months	age, gender, marital status, dummy for having children with less than 12 years in the household, race, schooling, dummies for having changed job, private sector, part time managerial position, supervisor, firm size, region and time.	Negative (-3% for OLS) but not statistically significant for FE.	Yes Fixed effects

Panel B: Papers using worker level data - developing countries

Name of study	Country and time period	Dependent variable	Training	Other independent variables	Effects training on wages	Controls endogeneity training
Chung (2000)	Worker Level Data for Malaysia (1976, 1988)	log hourly wage	dummy: having ever participated in training	age, marital status, nationality, schooling, dummies for employers and unpaid family workers.	20%-30% (for women)	Yes Excluded instruments (selection model): having a bank account, level of education in 1976, and parents occupational status.
Frazer (2006)	Worker Level Data for Ghana (1991-1999)	log hourly wage	dummy: having participated in an apprenticeship	gender, potential experience, schooling.	Not statistically different from zero for the whole sample but 17% for self-employed.	No
Johanson and Wanga (2008)	Worker Level Data for Tanzania (2006)	log hourly wage	dummy: having ever participated in training (per type: on-the-job, informal apprenticeship, vocational, college/advanced)	experience, gender, occupation, schooling, rural dummy and region.	38% for on-the-job training, 27% for formal apprenticeship, 47% for vocational training and 77% for college certificated training.	No

Table A1 Review of the literature on wage returns of job training (Continued)

Kahyarara and Teal (2008)	Worker Level Data for Tanzania (1997-2000)	log monthly wage	dummies: current and past on-the-job training and going on a short training course in the previous six months	gender, potential experience, occupation, schooling, tenure, dummy for capital city, firm fixed effects.	22% for current on-the-job training, not statistically different from zero for past on-the-job training and 17% for short training courses.	No
Monk, Sandefur and Teal (2008)	Worker Level Data for Ghana (1984, 2000)	log monthly wage	dummy: having participated in an apprenticeship	gender, potential experience, schooling, log hours worked per week, IQ score, interaction between apprenticeship and schooling, city.	50% for people with no formal education. The return declines as education rises.	Yes Members of the household that also made an apprenticeship, dummy for household access to credit, and a dummy for having internal piped water in the house as a wealth indicator.
Rosholm, Nielsen, Debalen (2007)	Matched employer-employee data for Kenya and Zambia (1995)	log monthly wage	dummy: having participated in training in the previous 12 months	age, ethnicity, experience, gender, occupation, schooling, tenure, union participation and familiar relations within the owners of the firm, ownership, industry, location, size, financial situation, skill demand, turnover, unionization, training annual expenses.	2.3% for Kenya and not statistically different from zero for Zambia.	Yes Matching Estimators Method (Local Linear Matching)
Almeida, Faria	Matched employer-employee data for Malaysia (2002) and Thailand (2004)	log hourly wage	dummy: having received formal on the job training since having joined the firm	educational attainment, gender, age, tenure, experience, marital status, occupation, union participation, computer, bank account, internet transaction, training at a previous employer, size, foreign capital, exports, education of the work force, education of the manager, new production technologies, industry, region.	7.7% for Malaysia and 4.5% for Thailand.	Yes Matching Estimators Method (Local Linear Matching)

Table A1 Review of the literature on wage returns of job training (Continued)

Panel C: Papers using firm level data						
Name of study	Data, country and time period	Dependent variable	Training	Other independent variables	Effects training on wages	Controls endogeneity training
Almeida and Carneiro (2008)	Firm Level Data for Portugal (1995-1999)	log value added per employee	average number of hours of training per employee	log employees, log capital stock, share occupation group, share low educated workers, share males workforce, cubic polynomial on average wage workforce, year dummies, region and sector	24% for firms providing training.	Yes First differences, GMM: past level of training as a instrument for current training.
Barrett and O'Connel (1999)	Firm Level Data for Ireland (1993, 1995)	productivity (out-put divided by total employment) growth	average training days per worker	investment, change on employment, sector, size, innovation, restructuring, management quality, dummies for labor incentives strategies.	Increasing one day of training per worker increases productivity growth by 0.03%.	Yes Dependent variable is productivity growth
Tan and Lopez-Acevedo (2003)	Firm Level Data for Mexico (1992, 1999)	log monthly wage	dummy: firm offered training in the previous 12 months	average years of schooling of the workforce, percentage of women, occupation, ownership, exports, size, industry and region	Training returns increased from 5% to 7% from 1992 to 1999.	Yes Excluded instruments (selection model): years in operation, R&D, computerization, unionization.
Panel D: Papers using industry level data						
Name of study	Country and time period	Dependent variable	Training	Other independent variables	Effects training on wages	Controls endogeneity training
Dearden, Reed and Reenen (2000)	Industry Level Data for U.K. (1984-1996)	log hourly wage	industry aggregated incidence for training in the previous 4 weeks	log capital per worker, log hours per worker, log of R&D over sales, region, time and tenure dummies, proportion of: men, age groups, occupation, qualified workers, small firms.	Raising training incidence by 5% increases wages and productivity by 1.6% and 4% respectively.	Yes Panel data: Within groups estimator

Table A2 Variable definitions

Variable	Definition
Firm training	Dummy variable equal to 1 if the firm offered formal training to its workers in the previous year.
Micro, small, medium, large, and very large	Dummy variables equal to 1 if the total number of employees in the firm is between 1 and 9, between 10 and 49, between 50 and 99, between 100 and 249 or greater than 250, respectively.
Sector	Two digit industries/services: Food Processing; Textiles; Garments; Wood Products/Furniture; Pharmaceuticals/Chemicals; Rubber and Plastics, Household Electrical Appliances; Electronics; Automobiles parts; Machinery and Equipment / Information Technology; Communication Services; Accounting and Related Services; Advertising and Marketing; Business Logistics.
Full foreign-owned	Dummy variable equal to 1 if 100% of the firm's capital is owned by foreigners.
Majority foreign-owned	Dummy variable equal to 1 if more than 50% and less than 100% of the firm's capital is owned by foreigners.
Minority foreign-owned	Dummy variable equal to 1 if more than 0% but less than 50% of the firm's capital is owned by foreigners.
Domestic	Dummy variable equal to 1 if 100% of the firm's capital is owned by domestic entities.
Exporter	Dummy variable equal to 1 if the firm exports are more than 10% of its sales
Share of skilled workers	Percentage of the firm's workforce that are managers, professionals or skilled production workers.
Education labor force	Average years of schooling of the workforce.
Age firm	Year of the survey minus the year when the firm started operations.
Education of manager	Dummy variable equal to 1 if the manager of the firm owns a degree.
New production technology	Dummy variable equal to 1 if the firm introduced in the previous two years a new technology that substantially changed production.
Training in the firm	Dummy variable equal to 1 if the worker received formal training since he joined the firm.
Hourly wage	Hourly wage in USD (exchange rate at the time of the survey)
Ln wage	Natural logarithm of hourly wage.
Woman	Dummy variable equal to 1 if the worker is a woman.
Age of worker	Age of the worker
Level of education	Dummy variables equal to 1 if the highest level of formal education of the worker is a degree, diploma, upper secondary, lower secondary, primary, respectively.
Vocational education	Dummy variable equal to 1 if the worker received vocational education.
Potential experience	Year of the survey minus the year when the worker found his first permanent job.
Tenure with the firm	Year of the survey minus the year when the worker joined the firm.
Single	Dummy variable equal to 1 if the worker is single
Native	Dummy variable equal to 1 if the worker is from the same country were the firm is registered.
Studied abroad	Dummy variable equal to 1 if the worker studied abroad.
Unionized	Dummy variable equal to 1 if the worker belongs to a trade union.
Occupation dummies	Dummy variables equal to 1 if the employee works as a manager, professional, skill production worker, unskilled production worker, non-production worker, apprentice, respectively.
Computer	Dummy variable equal to 1 if the worker owns a personal computer at home.
Bank account	Dummy variable equal to 1 if the worker has a bank account.
Transaction internet	Dummy variable equal to 1 if the worker has ever bought or sold a good through the internet.
Training other firm	Dummy variable equal to 1 if the worker received formal training at the previous employer.

Source: *Enterprise Surveys* (World Bank), Malaysia (2002) and Thailand (2004).

Table A3 Summary statistics of the main firm level characteristics

Variable	Malaysia		Thailand	
	Obs. (1)	Mean (2)	Obs. (3)	Mean (4)
Firm training	1,145	0.51	1,385	0.76
In-house training	551	0.30	1,006	0.46
Outside training	570	0.11	1,049	0.08
Share costs with employees	553	0.05	1,056	0.03
Size				
Micro	1,148	0.06	1,385	0.02
Small	1,148	0.43	1,385	0.28
Medium	1,148	0.20	1,385	0.19
Large	1,148	0.16	1,385	0.23
Very large	1,148	0.15	1,385	0.29
Ownership				
Full foreign-owned	1,128	0.13	1,384	0.07
Majority foreign-owned	1,128	0.07	1,384	0.06
Minority foreign-owned	1,128	0.10	1,384	0.13
Domestic	1,128	0.70	1,384	0.74
Exporter	941	0.67	1,385	0.62
Share of skilled workers	1,149	0.49	1,385	0.24
Education labor force	1,149	10.12	1,385	9.73
Sector				
Food processing	1,152	0.18	1,385	0.13
Textiles	1,152	0.03	1,385	0.13
Garments	1,152	0.09	1,385	0.12
Automobiles parts	1,152	0.03	1,385	0.10
Electronics	1,152	0.07	1,385	0.11
Rubber and plastics	1,152	0.22	1,385	0.06
Wood products/furniture	1,152	0.00	1,385	0.17
Machinery and equipment	1,152	0.07	1,385	0.09
Pharmaceuticals/chemicals	1,152	0.03	-	-
Household electrical appliances	1,152	0.08	-	-
Information technology	1,152	0.03	-	-
Communication services	1,152	0.01	-	-
Accounting and related services	1,152	0.09	-	-
Advertising and marketing	1,152	0.02	-	-
Business logistics	1,152	0.08	-	-

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: Table reports the means of variables listed in the table. Column (1) reports the means for the sample of firms in Malaysia, and column (2) reports the means for the sample of firms in Thailand.

Table A4 Summary statistics of the main worker level characteristics

Variable	Malaysia (1)	Thailand (2)
Training in the firm	0.33	0.52
Hourly wages	0.41	-0.14
Woman	0.44	0.55
Age of worker	34.18	32.41
Education		
Degree	0.08	0.20
Diploma	0.12	0.11
Upper secondary	0.36	0.20
Lower secondary	0.27	0.18
Primary	0.15	0.30
Illiterate	0.01	0.00
Vocational education	0.15	0.06
Potential experience	13.61	13.85
Tenure with the firm	7.24	5.55
Single	0.35	0.36
Studied abroad	0.07	0.01
Unionized	0.04	0.01
Occupation		
Managers	0.14	0.06
Professionals	0.08	0.08
Skill production	0.36	0.29
Unskilled production	0.24	0.37
Non-production	0.17	0.18
Apprentice	0.01	0.01
Observations	6,679	9,418

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: Table reports the means of variables listed in the table. Column (1) reports the means for the sample of workers in Malaysia, and column (2) reports the means for the sample of workers in Thailand. The wage variables are in USD at current prices at the time of the survey.

Table A5 Average marginal effects for Malaysian firms

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Small	0.182** [0.082]	0.158* [0.088]	0.166* [0.089]	0.151* [0.087]	0.182** [0.092]	0.180* [0.093]	0.167* [0.100]
Medium	0.410*** [0.063]	0.384*** [0.072]	0.404*** [0.072]	0.379*** [0.073]	0.376*** [0.076]	0.376*** [0.077]	0.373*** [0.085]
Large	0.516*** [0.043]	0.487*** [0.057]	0.503*** [0.056]	0.483*** [0.057]	0.462*** [0.062]	0.455*** [0.064]	0.475*** [0.068]
Very large	0.589*** [0.029]	0.574*** [0.040]	0.588*** [0.039]	0.573*** [0.040]	0.560*** [0.042]	0.549*** [0.046]	0.564*** [0.048]
Age firm	-0.005 [0.004]	-0.002 [0.003]	-0.001 [0.003]	-0.001 [0.003]	0.00 [0.003]	0.00 [0.003]	0.00 [0.003]
Exporter		0.243*** [0.045]	0.250*** [0.045]	0.244*** [0.045]	0.240*** [0.047]	0.228*** [0.048]	0.240*** [0.049]

Table A5 Average marginal effects for Malaysian firms (Continued)

Minority foreign ownership	0.052	0.043	0.04	-0.002	-0.005	0.011	
	[0.065]	[0.065]	[0.066]	[0.068]	[0.069]	[0.071]	
Majority foreign ownership	0.059	0.047	0.049	0.028	0.024	0.043	
	[0.077]	[0.079]	[0.078]	[0.079]	[0.078]	[0.079]	
Full foreign ownership	0.131**	0.121**	0.121**	0.100	0.107*	0.113*	
	[0.058]	[0.059]	[0.059]	[0.062]	[0.064]	[0.067]	
Share of skilled workers		0.180**	-	-	-	-	
		[0.071]					
Education labor force			0.014***	0.011**	0.010**	0.008*	
			[0.005]	[0.005]	[0.005]	[0.005]	
Education of manager				0.228***	0.234***	0.234***	
				[0.042]	[0.042]	[0.043]	
New production technology					0.136***	0.126***	
					[0.045]	[0.047]	
Industry fixed effects?	Yes	Yes	Yes	Yes	Yes	Yes	No
Region fixed effects?	Yes	Yes	Yes	Yes	Yes	Yes	No
Industry-region fixed effects?	No	No	No	No	No	No	Yes
Observations	1,132	910	909	909	890	887	869

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

The dependent variable is a dummy variable that assumes the value 1 if the firm offered formal on-the-job training to its employees. The table reports the marginal effects (at mean values) on the firm's propensity to train from probit regressions. Robust standard errors are in brackets. * significant at 10%, ** significant at 5%, *** significant at 1%. All variables are defined in Table A2. Micro firms (with less than 10 employees) is the omitted size group. Age squared is also included in the regressions (not reported). Industry fixed effects refer to 2 digit industry or service.

Table A6 Average marginal effects for Thai Firms

	(1)	(2)	(3)	(4)	(5)	(6)	(7)
Small	0.055	0.045	0.041	0.079	0.077	0.078	0.072
	[0.062]	[0.062]	[0.063]	[0.058]	[0.059]	[0.057]	[0.060]
Medium	0.183***	0.169***	0.166***	0.185***	0.182***	0.177***	0.170***
	[0.037]	[0.039]	[0.039]	[0.035]	[0.036]	[0.035]	[0.039]
Large	0.253***	0.235***	0.232***	0.251***	0.248***	0.240***	0.238***
	[0.034]	[0.035]	[0.035]	[0.033]	[0.033]	[0.033]	[0.035]
Very large	0.326***	0.299***	0.295***	0.318***	0.310***	0.296***	0.279***
	[0.039]	[0.040]	[0.041]	[0.039]	[0.040]	[0.040]	[0.041]
Age firm	0.004	0.004	0.004	0.004	0.005	0.005	0.006
	[0.005]	[0.004]	[0.004]	[0.004]	[0.004]	[0.004]	[0.005]
Exporter		0.059**	0.060**	0.049*	0.045	0.042	0.048*
		[0.028]	[0.028]	[0.027]	[0.027]	[0.027]	[0.027]
Minority foreign ownership		0.004	0.004	-0.006	-0.008	-0.009	-0.002
		[0.036]	[0.036]	[0.037]	[0.037]	[0.038]	[0.038]
Majority foreign ownership		0.123***	0.123***	0.105**	0.104**	0.099**	0.117***
		[0.040]	[0.040]	[0.045]	[0.046]	[0.048]	[0.040]
Full foreign ownership		-0.031	-0.03	-0.062	-0.069	-0.073	-0.042
		[0.067]	[0.067]	[0.069]	[0.070]	[0.072]	[0.064]
Share of skilled workers			-0.035	-	-	-	-
			[0.049]				

Table A6 Average marginal effects for Thai Firms (Continued)

Education labor force				0.026***	0.024***	0.022***	0.021***
				[0.005]	[0.005]	[0.005]	[0.005]
Education of manager					0.059**	0.051**	0.071***
					[0.026]	[0.026]	[0.026]
New production technology						0.078***	0.085***
						[0.023]	[0.023]
Industry fixed effects?	Yes	Yes	Yes	Yes	Yes	Yes	No
Region fixed effects?	Yes	Yes	Yes	Yes	Yes	Yes	No
Industry-region fixed effects?	No	No	No	No	No	No	Yes
Observations	1,278	1,277	1,277	1,277	1,277	1,277	1,348

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

The dependent variable is a dummy variable that assumes the value 1 if the firm offered formal on-the-job training to its employees. The table reports the marginal effects (at mean values) on the firm's propensity to train from probit regressions. Robust standard errors are in brackets. * significant at 10%, ** significant at 5%, *** significant at 1%. All variables are defined in Table A2. Micro firms (with less than 10 employees) is the omitted size group. Age squared is also included in the regressions (not reported). Industry fixed effects refer to 2 digit industry or service.

Table A7 Average marginal effects for Malaysian workers

	(1)	(2)	(3)	(4)	(5)
Degree	0.713***	0.653***	0.649***	0.606***	0.523***
	[0.181]	[0.182]	[0.184]	[0.184]	[0.186]
Diploma	0.783***	0.730***	0.726***	0.711***	0.608***
	[0.176]	[0.177]	[0.179]	[0.179]	[0.180]
Upper secondary	0.552***	0.521***	0.516***	0.514***	0.458***
	[0.170]	[0.170]	[0.171]	[0.171]	[0.172]
Lower secondary	0.308*	0.287*	0.283*	0.282	0.266
	[0.170]	[0.170]	[0.172]	[0.171]	[0.172]
Primary	0.026	0.018	0.015	0.015	0.036
	[0.175]	[0.175]	[0.175]	[0.175]	[0.176]
Vocational education	0.170***	0.162***	0.162***	0.162***	0.101**
	[0.048]	[0.048]	[0.049]	[0.049]	[0.049]
Woman	-0.095***	-0.098***	-0.099***	-0.095***	-0.071
	[0.037]	[0.037]	[0.037]	[0.037]	[0.038]
Age of worker	0.011	0.011	0.011	0.012	0.020
	[0.019]	[0.019]	[0.019]	[0.019]	[0.019]
Tenure with the firm	0.047***	0.046***	0.046***	0.046***	0.069***
	[0.009]	[0.009]	[0.009]	[0.009]	[0.001]
Potential experience	-0.011	-0.011	-0.012	-0.012	-0.032***
	[0.009]	[0.009]	[0.009]	[0.009]	[0.009]
Single	-0.180***	-0.182***	-0.182***	-0.181***	-0.146***
	[0.044]	[0.044]	[0.044]	[0.045]	[0.045]
Manager	0.142**	0.128**	0.129**	0.117**	0.060
	[0.064]	[0.064]	[0.064]	[0.064]	[0.066]
Professional	0.165**	0.157	0.157	0.153	0.118
	[0.076]	[0.076]	[0.076]	[0.076]	[0.078]

Table A7 Average marginal effects for Malaysian workers (Continued)

Skill Production	-0.031 [0.053]	-0.021 [0.053]	-0.021 [0.053]	-0.018 [0.053]	-0.017 [0.054]
Unskilled Production	-0.298*** [0.061]	-0.280*** [0.061]	-0.279*** [0.061]	-0.276*** [0.061]	-0.269*** [0.062]
Apprentice	0.433 [0.203]	0.453 [0.203]	0.455 [0.203]	0.462 [0.203]	0.506** [0.204]
Unionized	0.522*** [0.087]	0.519*** [0.087]	0.519*** [0.088]	0.514*** [0.088]	0.531*** [0.088]
Computer		0.119*** [0.038]	0.119*** [0.039]	0.112*** [0.039]	0.108*** [0.040]
Bank Account			0.019 [0.095]	0.016 [0.095]	-0.040 [0.095]
Transaction Internet				0.334*** [0.094]	0.277*** [0.095]
Training at the Previous Firm					0.793*** [0.048]
Observations	6,679	6,679	6,679	6,679	6,679

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

The dependent variable is a dummy variable that assumes the value 1 if the received formal on-the-job training since he joined the firm. The table reports the marginal effects (at mean values) on the worker's propensity to be trained from probit regressions. Robust standard errors are in brackets. * significant at 10%, ** significant at 5%, *** significant at 1%. All variables are defined in Table A2. Firm Base Specification is the same as in column (7) of Table A5. Illiterate and Non-production workers are the omitted education and occupation groups. Age squared, Tenure squared and Experience squared are also included in the regressions (not reported).

Table A8 Average marginal effects for Thai workers

	(1)	(2)	(3)	(4)	(5)
Degree	0.797*** [0.210]	0.717*** [0.211]	0.593*** [0.216]	0.586*** [0.216]	0.426** [0.215]
Diploma	0.662*** [0.210]	0.617*** [0.211]	0.494*** [0.215]	0.493*** [0.216]	0.353 [0.215]
Upper secondary	0.484*** [0.207]	0.463*** [0.208]	0.347* [0.212]	0.345 [0.212]	0.226 [0.212]
Lower secondary	0.340 [0.207]	0.334 [0.208]	0.239 [0.212]	0.237 [0.212]	0.114 [0.211]
Primary	0.116 [0.205]	0.119 [0.206]	0.051 [0.210]	0.049 [0.210]	-0.034 [0.209]
Vocational education	0.075 [0.058]	0.067 [0.058]	0.058 [0.058]	0.050 [0.058]	-0.015 [0.059]
Woman	-0.048* [0.029]	-0.046* [0.029]	-0.066** [0.029]	-0.066** [0.029]	-0.072* [0.030]
Age of worker	0.025 [0.015]	0.026* [0.015]	0.023 [0.016]	0.023 [0.016]	0.020 [0.016]
Tenure with the firm	0.085*** [0.007]	0.085*** [0.008]	0.083*** [0.008]	0.082*** [0.008]	0.100*** [0.008]
Potential experience	0.005 [0.007]	0.004 [0.007]	0.003 [0.007]	0.003 [0.007]	-0.003 [0.008]

Table A8 Average marginal effects for Thai workers (Continued)

Single	0.001 [0.032]	0.003 [0.032]	0.012 [0.033]	0.012 [0.033]	0.022 [0.033]
Manager	-0.002 [0.068]	-0.042 [0.068]	-0.047 [0.068]	-0.065 [0.068]	-0.087 [0.070]
Professional	-0.021 [0.059]	-0.041 [0.059]	-0.047 [0.059]	-0.051 [0.059]	-0.055 [0.059]
Skill production	0.176*** [0.045]	0.183*** [0.045]	0.185*** [0.045]	0.185*** [0.045]	0.181*** [0.046]
Unskilled production	0.218*** [0.046]	0.229*** [0.046]	0.239*** [0.046]	0.237*** [0.046]	0.225*** [0.047]
Apprentice	-0.291 [0.173]	-0.281 [0.173]	-0.272 [0.173]	-0.272 [0.173]	-0.234 [0.174]
Unionized	0.635*** [0.143]	0.636*** [0.143]	0.635*** [0.143]	0.637*** [0.143]	0.628*** [0.144]
Computer		0.177*** [0.039]	0.163*** [0.039]	0.158*** [0.039]	0.143*** [0.040]
Bank account			0.465*** [0.052]	0.465*** [0.052]	0.475*** [0.052]
Transaction internet				0.293** [0.140]	0.193 [0.141]
Training at the previous firm					0.607*** [0.014]
Observations	9,418	9,418	9,418	9,418	9,418

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

The dependent variable is a dummy variable that assumes the value 1 if the received formal on-the-job training since he joined the firm. The table reports the marginal effects (at mean values) on the worker's propensity to be trained from probit regressions. Robust standard errors are in brackets. * significant at 10%, ** significant at 5%, *** significant at 1%. All variables are defined in Table A2. Firm Base Specification is the same as in column (7) of Table A6. Illiterate and Non-production workers are the omitted education and occupation groups. Age squared, Tenure squared and Experience squared are also included in the regressions (not reported).

Table A9 Wage returns to job training in Malaysia

	(1)	(2)	(3)
Training in the firm	0.431*** [0.037]	0.081*** [0.026]	0.043*** [0.027]
Degree		0.739*** [0.077]	0.637*** [0.071]
Diploma		0.453*** [0.073]	0.383*** [0.067]
Upper secondary		0.209*** [0.067]	0.187*** [0.061]
Lower secondary		0.181*** [0.066]	0.152*** [0.060]
Primary		0.010 [0.063]	0.007 [0.058]
Vocational education		0.114*** [0.023]	0.075*** [0.020]

Table A9 Wage returns to job training in Malaysia (Continued)

Woman		-0.203***	-0.187***
		[0.019]	[0.018]
Age of worker		0.031***	0.029***
		[0.008]	[0.008]
Tenure with the firm		0.027***	0.027***
		[0.005]	[0.004]
Potential experience		0.017***	0.016***
		[0.004]	[0.004]
Single		-0.072***	-0.067***
		[0.020]	[0.019]
Manager		0.384***	0.381***
		[0.036]	[0.032]
Professional		0.302***	0.264***
		[0.041]	[0.037]
Skill production		-0.079***	-0.091***
		[0.033]	[0.028]
Unskilled production		-0.350***	-0.366***
		[0.035]	[0.031]
Apprentice		-0.276***	-0.300***
		[0.068]	[0.072]
Unionized		-0.064**	-0.085**
		[0.055]	[0.051]
Computer		0.173***	0.153***
		[0.020]	[0.018]
Bank account		0.179***	0.130***
		[0.041]	[0.041]
Transaction internet		0.175***	0.094**
		[0.062]	[0.048]
Training at the previous firm		0.091***	0.068***
		[0.026]	[0.026]
Firm base specification	No	No	Yes
Observations	6,679	6,679	6,679

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: The table reports the estimates for equation 5 in the text. Robust standard errors, clustered at firm level, are in brackets. * significant at 10%, ** significant at 5%, *** significant at 1%. The dependent variable is the log hourly wage. Column (1) only controls for job training, column (2) adds baseline worker characteristics as in column (8) of Table A7 and A8. Column (3) adds firm baseline characteristics as in column (7) of Tables A5 and A6. Illiterate and Non-production workers are the omitted education and occupation groups. Age squared, Tenure squared and Experience squared are also included in the regressions (not reported).

Table A10 Wage returns to job training in Thailand

	(1)	(2)	(3)
Training in the firm	0.284*** [0.022]	0.062*** [0.014]	0.042*** [0.015]
Degree		0.752*** [0.065]	0.743*** [0.069]
Diploma		0.477*** [0.063]	0.454*** [0.068]
Upper secondary		0.295*** [0.061]	0.283*** [0.066]
Lower secondary		0.176*** [0.060]	0.178** [0.065]
Primary		0.056 [0.059]	0.087 [0.064]
Vocational education		-0.012 [0.024]	-0.030 [0.024]
Woman		-0.168*** [0.013]	-0.163*** [0.012]
Age of worker		0.026*** [0.007]	0.022*** [0.006]
Tenure with the firm		0.035*** [0.004]	0.032*** [0.004]
Potential experience		0.009*** [0.001]	0.008*** [0.001]
Single		-0.031** [0.013]	-0.046*** [0.012]
Manager		0.568*** [0.041]	0.585*** [0.041]
Professional		0.198*** [0.028]	0.190*** [0.027]
Skill production		-0.025 [0.018]	-0.020 [0.019]
Unskilled production		-0.180*** [0.018]	-0.179*** [0.019]
Apprentice		-0.081 [0.124]	-0.186** [0.125]
Unionized		0.028 [0.056]	-0.079 [0.053]
Computer		0.214*** [0.018]	0.189*** [0.017]
Bank account		0.111*** [0.019]	0.087*** [0.019]

Table A10 Wage returns to job training in Thailand (Continued)

Transaction internet		0.312***	0.281***
		[0.064]	[0.060]
Training at the previous firm		0.011	0.027**
		[0.016]	[0.015]
Firm base specification	No	No	Yes
Observations	9418	9418	9418

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: The table reports the estimates for equation 5 in the text. Robust standard errors, clustered at firm level, are in brackets. * significant at 10%, ** significant at 5%, *** significant at 1%. The dependent variable is the log hourly wage. Column (1) only controls for job training, column (2) adds baseline worker characteristics as in column (8) of Table A7 and A8. Column (3) adds firm baseline characteristics as in column (7) of Tables A5 and A6. Illiterate and Non-production workers are the omitted education and occupation groups. Age squared, Tenure squared and Experience squared are also included in the regressions (not reported).

Table A11 Balancing tests

	Malaysia			Thailand		
	Treated means	Matched means	T-test	Treated means	Matched means	T-test
	(1)	(2)	(3)	(4)	(5)	(6)
Training other firm	0.323	0.332	-0.62	0.322	0.310	1.27
Woman	0.424	0.436	-0.17	0.558	0.554	0.41
Age of worker	33.396	32.930	1.88	32.610	32.605	0.04
Education						
Degree	0.137	0.136	0.02	0.257	0.253	0.47
Diploma	0.193	0.200	-0.61	0.134	0.139	-0.69
Upper secondary	0.416	0.408	0.54	0.216	0.220	-0.50
Lower secondary	0.197	0.193	0.34	0.170	0.163	0.84
Primary	0.052	0.055	-0.48	0.220	0.221	-0.15
Illiterate	0.001	0.003	-0.77	0.003	0.003	-0.18
Vocational education	0.237	0.259	-1.72	0.070	0.072	-0.40
Potential experience	12.654	12.140	2.10	14.021	13.802	0.80
Tenure with the firm	7.566	6.996	3.14	6.107	6.055	0.52
Single	0.317	0.342	-1.77	0.366	0.368	-0.21
Studied abroad	0.079	0.073	0.71	0.011	0.012	-0.70
Unionized	0.075	0.070	0.63	0.021	0.018	1.47
Occupation						
Managers	0.208	0.204	0.31	0.075	0.078	-0.52
Professionals	0.134	0.146	-1.19	0.098	0.094	0.70
Skill production	0.344	0.339	0.30	0.300	0.295	0.58
Unskilled production	0.119	0.122	-0.31	0.332	0.320	1.17
Non-production	0.186	0.177	0.76	0.190	0.207	-2.12
Apprentice	0.010	0.011	-0.44	0.004	0.005	-0.60

Source: Authors' calculations based on the *Enterprise Surveys* (World Bank).

Note: The table reports balancing tests between the sample means of the variables listed. We contrast the means of the subsample of treated and untreated individuals. The t-test reported in column (3) and (6) for Malaysia and Thailand, respectively, verifies if the difference between the means of the variables reported is, for each country, statistically different from zero across the two samples. Treated individuals are those that participated in training, and untreated individuals are those reporting not having participated in training.

A2. Least squares returns to on-the-job training

Following Mincer, 1974, we assume that (log) wages are a linear function of several human capital and other worker characteristics, and of firm characteristics:

$$\ln w_{ij} = \beta Train_{ij} + \lambda X_{ij} + \phi Z_j + \varepsilon_{ij} \quad (5)$$

where w_{ij} is the worker's hourly wage (local currency) for worker i in firm j , $Train_{ij}$ is a dummy variable equal to one if the worker received formal training since he joined firm j , X_{ij} is a vector of the worker's characteristics, Z_j is a vector of firm level characteristics and ε_{ij} captures the unobserved characteristics of worker i in firm j correlated with hourly wages. Our main parameter of interest is the coefficient β . β captures the percentage point difference in the hourly wage for workers who have received formal on-the-job training in firm j . We estimate equation 1 with least squares and cluster the standard errors at the firm level.

The least squares estimates for β are consistent if $Train_{ij}$ is uncorrelated with the error term ε_{ij} . However, this assumption may not hold. On the one hand, there is likely self selection into on-the-job training. We have shown that workers with certain observable characteristics (and most likely also unobservable) are more likely to have taken on-the-job training programs than others. Therefore, it is possible that the higher earnings for those who are trained are caused not by training itself but because those taking up training could have a greater earning capacity and ability than the non-trainees. In this case, the least squares estimates of β will probably be upward biased due to a possible "ability bias". On the other hand, if the variable on-the-job training is measured with error, the least squares estimates could be downward biased. Therefore the overall sign of the least square bias is unclear.

We minimize the first problem by accounting in the reduced form for several observable individual and firms characteristics simultaneously correlated with training and also with hourly wages. In particular, in X_{ij} we include detailed information on schooling, gender, age, tenure in the firm, potential experience, marital status, occupation, ethnicity and age. In Z_j we include information on firm size, foreign ownership, exports, average schooling of the workforce, managerial ability, degree of technological innovation, industry and geographical location of the firm.

Tables A9 and A10 report the least square estimates for β when exploring different specifications and after clustering the standard errors at the firm level. Column (1) controls only for training incidence since joining the firm, column (2) adds the baseline worker characteristics (as in column (6) of Tables A7 and A8)²¹, column (3) adds the baseline firm characteristics (reported in column (7) of Tables A5 and A6)²² to the specification in column (3).

The OLS estimates strongly suggest that there are positive returns to the investment in on-the-job training in both countries. As expected, the magnitude of the returns decreases as we introduce additional firm and worker controls. In Table A9, the wage returns of on-the-job training for Malaysia start at 43.1% but fall to 8.1% when we control for worker characteristics and to 4.3% once we control for firm characteristics. Table A10 reports similar findings for Thailand. Returns start at 28.4% falling to 6.2% when we include workers characteristics and to 4.2% when we include firm characteristics²³.

Competing interests

The "Journal of Labor & Development" is committed to the IZA Guiding Principles of Research Integrity. The authors declare that they have observed these principles.

Authors' contributions

Rita Almeida and Marta Faria carried out the treatment and econometric analysis of the data, and they both drafted the manuscript. The authors read and approved the final manuscript.

Acknowledgement

We thank the anonymous comments of a referee and the editor as well as the financial support provided by the Development Economics Research Support Grant (World Bank). Address: Travessa Palma, 1649-023 Lisboa. E-mail: martalincefaria@gmail.com, ralmeida@worldbank.org. The findings expressed in this paper are those of the author and do not necessarily represent the views of the World Bank. We also thank Leonor Modesto, Miguel Gouveia, Pedro Raposo and participants of the IZA conference, Cape Town 3rd and 4th of May 2010. Responsible editor: Hartmut Lehmann.

Author details

¹The World Bank and IZA, Lisbon, Portugal. ²Católica Lisbon, SBE, Lisbon, Portugal.

Received: 31 January 2014 Accepted: 12 June 2014

Published: 10 Nov 2014

References

- Acemoglu D, Pischke J: **The Structure of Wages and Investment in General Training.** *J Polit Econ* 1999, **107**(3):539–572.
- Aghion P, Howitt P: *Endogenous Growth Theory.* Cambridge, MA: MIT press; 1998. MIT Press.
- Albert C, García Serrano C, Hernanz V: **On-the-job training in Europe: Determinants and wage returns.** *Int Lab Rev* 2010, **149**(3):315–341.
- Almeida R: **Innovation and Openness in East Asia: Have they Increased the Demand for Educated Workers?** *Asia Pac Dev J* 2010, **17**(1):63–95.
- Almeida RK, Aterido R: *Investment in job training: why are SMES lagging so much behind?* Policy Research Working Paper Series 5358. Washington, D.C: The World Bank; 2010.
- Almeida R, Aterido R: **On-the-Job Training and Rigidity of Employment Protection in the Developing World: Evidence from Differential Enforcement.** *Labour Economics* 2011, **18**:S71–S82.
- Almeida R, Carneiro P: **The return to the firm investment in human capital.** *Lab Econ* 2008, **16**(2009):97–106. Vol. 44, Issue 5, Pages 701–727.
- Almeida R, Cho Y: **On-the-job training: Stylized facts and policies around the developing world in Almeida.** In *The Right skills for the job?*. Edited by Berhman and Robalino. Washington, D.C: Rethinking training policies in developing countries, World Bank Human Development Perspectives Series, The World Bank; 2012.
- Ariga K, Brunello G: *Are the More Educated Receiving More Training?* Evidence from Thailand. Bonn: IZA Discussion Papers 577, Institute for the Study of Labor (IZA); 2002.
- Barret A, O'Connell P: *Does Training Generally Work? The Returns to In-Company Training.* Bonn: Institute for the Study of Labor (IZA); 1999. IZA Discussion Paper No. 51, August 1999.
- Bassinini A, Booth A, Brunello G, De Paola M, Leuven E: *Workplace Training in Europe.* Bonn: Institute for the Study of Labor (IZA); 2005. IZA DP No. 1640.
- Becker G: *Human Capital.* Chicago: The University of Chicago Press; 1964.
- Budria S, Pereira P: *On the Returns to Training in Portugal.* Bonn: Institute for the Study of Labor (IZA); 2004. IZA DP No. 1429.
- Chung T: *The Returns to Education and Training: Evidence from the Malaysian Family Life Surveys.* Kent: University of Kent; 2000. Discussion Paper No. 00/07, University of Kent.
- Dearden L, Reed H, Reenen JV: *Who Gains When Workers Train? Training and Corporate productivity in a Panel of British Industries.* London, UK: The Institute for Fiscal Studies; 2000. WP 00/04.
- Dearden L, Reed H, Reenen JV: *The Impact of Training on Productivity and Wages: Evidence from British Panel Data,* Volume 4. 68th edition. Oxford: Oxford Bulletin of Economics and Statistics, Department of Economics, University of Oxford; 2006:397–421. 08.
- Fitzenberger B, Völter R: **Long-run effects of training programs for the unemployed in East Germany.** *Lab Econ* 2007, **14**(4):730–755.
- Frazer G: **Learning the Master's Trade: Apprenticeship and Human Capital in Ghana.** *J Dev Econ* 2006, **81**(2):259–298.
- Haelermans C, Borghans L: **Wage Effects of On-the-Job Training: A Meta-Analysis.** *Br J Ind Relat* 2012, **50**:502–528. 3 September 2012 0007–1080.
- Heckman J, Lalonde RJ, Smith JA: **The economics and econometrics of active labour market programs.** In *Handbook of Labor Economics.* Edited by Ashenfelter, Card. North Holland, Amsterdam: New York; 1999:1865–2097.
- Johanson R, Wanga G: *Skills Training for the Informal Sector in Tanzania.* Washington, D.C: World Bank Report; 2008.
- Kahyarara G, Teal F: **The Returns to Vocational Training and Academic Education: Evidence from Tanzania.** *World Dev* 2008, **36**(11):2223–2242.
- Leuven E, Oosterbeek H: **Firm-specific human capital as a shared investment: Comment.** *Am Econ Rev* 2001, **91**:342–347.
- Leuven E, Oosterbeek H: *A New Approach to Estimate the Wage Returns to Work-Related Training.* Bonn: Institute for the Study of Labor (IZA); 2002. IZA DP No. 526.
- Leuven E, Oosterbeek H: **Evaluating the Effect of Tax deductions on training.** *J Labor Econ* 2004, **22**(2):461–488. University of Chicago Press.

- Lillard L, Tan H: *Private Sector Training: Who Gets Its and What are its Effects?*. Santa Monica: RAND; 1986. RAND Report, R-3331-DOL.
- Lopez-Acevedo G, Tan H: *Mexico:in-firm Training for the Knowledge Economy*. Washington, D.C: World Bank Policy Research Working Paper 2957; 2003.
- Lucas R Jr: **On the mechanics of economic development**. *J Monet Econ* 1988, **22**(1):3–42. Elsevier.
- Monk C, Sandefur J, Teal F: *Does Doing an Apprenticeship Pay Off? Evidence from Ghana*. Oxford: University of Oxford working paper CSAE WPS/2008-08; 2008.
- Psacharopoulos G, Patrinos H: **Returns to investment in education: a further update**. *Educ Econ* 2004, **12**:111–134.
- Romer P: **Endogenous Technological Change**. *Journal of Political Economy* 1990, **98**(5):S71–102. University of Chicago Press.
- Rosenbaum PR, Rubin DB: **The central role of the propensity score in observational studies for casual effects**. *Biometrika* 1983, **70**(1):41–55.
- Rosholm M, Nielsen H, Dabalen A: **Evaluation of training in African enterprises**. *J Dev Econ* 2007, **84**(2007):310–329.
- Sousounis P: *The Impact of Work-Related Training on Employee Earnings: Evidence from Great Britain*. Munich; 2009. MPRA paper No. 14262.

10.1186/2193-9020-3-19

Cite this article as: Almeida and Faria: **The wage returns to on-the-job training: evidence from matched employer-employee data**. *IZA Journal of Labor & Development* 2014, **3**:19

Submit your manuscript to a SpringerOpen[®] journal and benefit from:

- ▶ Convenient online submission
- ▶ Rigorous peer review
- ▶ Immediate publication on acceptance
- ▶ Open access: articles freely available online
- ▶ High visibility within the field
- ▶ Retaining the copyright to your article

Submit your next manuscript at ▶ springeropen.com
