## ECONSTOR

Make Your Publications Visible.

## A Service of

 2BWLeibniz-Informationszentrum Wirtschaft Leibniz Information Centre for Economics

# Working Paper <br> Market Power and Welfare in Asymmetric Divisible Good Auctions 

CESifo Working Paper, No. 6261

## Provided in Cooperation with:

Ifo Institute - Leibniz Institute for Economic Research at the University of Munich

Suggested Citation: Manzano, Carolina; Vives, Xavier (2016) : Market Power and Welfare in Asymmetric Divisible Good Auctions, CESifo Working Paper, No. 6261, Center for Economic Studies and ifo Institute (CESifo), Munich

This Version is available at:
https://hdl.handle.net/10419/149348

## Standard-Nutzungsbedingungen:

Die Dokumente auf EconStor dürfen zu eigenen wissenschaftlichen Zwecken und zum Privatgebrauch gespeichert und kopiert werden.

Sie dürfen die Dokumente nicht für öffentliche oder kommerzielle Zwecke vervielfältigen, öffentlich ausstellen, öffentlich zugänglich machen, vertreiben oder anderweitig nutzen.

Sofern die Verfasser die Dokumente unter Open-Content-Lizenzen (insbesondere CC-Lizenzen) zur Verfügung gestellt haben sollten, gelten abweichend von diesen Nutzungsbedingungen die in der dort genannten Lizenz gewährten Nutzungsrechte.

[^0]

# CESifo Working Papers <br> www.cesifo.org/wp 

# Market Power and Welfare in Asymmetric Divisible Good Auctions 

Carolina Manzano Xavier Vives

CESifo Working Paper No. 6261<br>CATEGORY 11: Industrial Organisation<br>December 2016

An electronic version of the paper may be downloaded

- from the SSRN website: www.SSRN.com
- from the RePEc website: www.RePEc.org
- from the CESifo website: www.CESifo-group.org/wp


# Market Power and Welfare in Asymmetric Divisible Good Auctions 


#### Abstract

We analyze a divisible good uniform-price auction that features two groups each with a finite number of identical bidders. Equilibrium is unique, and the relative market power of a group increases with the precision of its private information but declines with its transaction costs. In line with empirical evidence, we find that an increase in transaction costs and/or a decrease in the precision of a bidding group.s information induces a strategic response from the other group, which thereafter attenuates its response to both private information and prices. A "stronger" bidding group -which has more precise private information, faces lower transaction costs, and is more oligopsonistic- has more market power and so will behave competitively only if it receives a higher per capita subsidy rate. When the strong group values the asset no less than the weak group, the expected deadweight loss increases with the quantity auctioned and also with the degree of payoff asymmetries. Market power and the deadweight loss may be negatively associated.


JEL-Codes: D440, D820, G140, E580.
Keywords: demand/supply schedule competition, private information, liquidity auctions, treasury auctions, electricity auctions.

Carolina Manzano*<br>Universitat Rovira i Virgili<br>Reus / Spain<br>carolina.manzano@urv.cat

Xavier Vives**<br>IESE Business School<br>University of Navarra<br>Barcelona / Spain<br>xvives@iese.edu

## December 2016

For helpful comments we are grateful to Roberto Burguet, Vitali Gretschko, Jacub Kastl, Leslie Marx, Meg Meyer, Antonio Miralles, Stephen Morris, Andrea Prat, and Tomasz Sadzik as well as seminar participants at the BGSE Summer Forum, Chicago, Columbia, EARIE, ESSET, Federal Reserve Board, Jornadas de Economía Industrial, Northwestern, Princeton, UPF, Queen Mary Theory Workshop and Wharton. We are also indebted to Jorge Paz for excellent research assistance.
*Financial support from project ECO2013-42884-P is gratefully acknowledged.
**Financial support from the Spanish Ministry of Economy and Competitiveness (Ref. ECO2015-63711P) and from the Generalitat de Catalunya, AGAUR grant 2014 SGR 1496, is gratefully acknowledged.

## 1 Introduction

Divisible good auctions are common in many markets, including government bonds, liquidity (refinancing operations), electricity, and emission markets. ${ }^{1}$ In those auctions, both market power and asymmetries among the participants are important; asymmetries can make market power relevant even in large markets. However, theoretical work in this area has been hampered by the difficulties of dealing with bidders that are asymmetric, have market power, and are competing in terms of demand or supply schedules in the presence of private information. This paper helps to fill that research gap by analyzing asymmetric uniform-price auctions in which there are two groups of bidders. Our aims are to characterize the equilibrium, to perform comparative statics and welfare analysis (from the standpoint of revenue and deadweight loss), and finally to derive implications for policy.

Divisible good auctions are typically populated by heterogenous participants in a concentrated market, and often we can distinguish a core group of bidders together with a fringe. The former are strong in the sense that they have better information, endure lower transaction costs, and are more oligopolistic (or oligopsonistic) than members of the fringe. As examples we discuss Treasury and liquidity auctions in addition to wholesale electricity markets.

Treasury auctions have bidders with significant market shares. That will be the case in most systems featuring a primary dealership, where participation is limited to a fixed number of bidders (this occurs, for example, in 29 out of 39 countries surveyed by Arnone and Iden 2003). A prime example are US Treasury auctions, which are uniform-price auctions. ${ }^{2}$ In these auctions, the top five bidders typically purchase close to half of US Treasury issues (Malvey and Archibald 1998). Armantier and Sbaï (2006) test for whether the bidders in French Treasury auctions are symmetric; these authors conclude that such auction participants can be divided into two distinct groups as a function of (a) their level of risk aversion and (b) the quality of their information about the value of the security to be sold. One small group consists of large financial institutions, which possess better information and are willing to take more risks. Kastl (2011) also finds evidence of two distinct groups of bidders in (uniform-price) Czech Treasury auctions. Other papers that report asymmetries between bidders in Treasury auctions include, among others, Umlauf (1993) for Mexico, Bjonnes (2001) for Norway, and Hortaçsu and McAdams (2010) for Turkey.

[^1]Bindseil et al. (2009) and Cassola et al. (2013) find that the heterogeneity of bidders in liquidity auctions is relevant. Cassola et al. (2013) analyze the evolution of bidding data from the European Central Bank's weekly refinancing operations before and during the early part of the financial crisis. The authors show that effects of the 2007 subprime market crisis were heterogeneous among European banks, and they conclude that the significant shift in bidding behavior after 9 August 2007 may reflect a change in the cost of short-term funding on the interbank market and/or a strategic response to other bidders. In particular, Cassola et al. (2013) find that one third of bidders experienced no change in their costs of short-term funds from alternative sources; this means that their altered bidding behavior was mainly strategic: bids were increased as a best response to the higher bids of rivals. ${ }^{3}$

Concentration is high also in other markets, such as wholesale electricity. This issue has attracted attention from academics and policy makers alike. A number of empirical studies have concluded that sellers have exercised significant market power in wholesale electricity markets (see e.g. Green and Newbery 1992; Wolfram 1998; Borenstein et al. 2002; Joskow and Kahn 2002). ${ }^{4}$ Most wholesale electricity markets prefer using a uniform-price auction to using a pay-as-bid auction (Cramton and Stoft 2006, 2007). In several of these markets (e.g., California, Australia), generating companies bid to sell power and wholesale customers bid to buy power. In such markets, asymmetries are prevalent. For example, some generators of wholesale electricity rely heavily on nuclear technology, which has flat marginal costs, whereas others rely mostly on fuel technologies, which have steep marginal costs. Holmberg and Wolak (2015) argue that, in wholesale electricity markets, information on suppliers' production costs is asymmetric. For evidence on the effect of cost heterogeneity on bidding in wholesale electricity markets, see Crawford et al. (2007) and Bustos-Salvagno (2015).

Our paper makes progress within the linear-Gaussian family of models by incorporating bidders' asymmetries with regard to payoffs and information. We model a uniform-price auction where asymmetric strategic bidders compete in terms of demand schedules for an inelastic supply (we can easily accommodate supply schedule competition for an inelastic demand). Bidders may differ in their valuations, transaction costs, and/or the precision of their private information. ${ }^{5}$

[^2]For simplicity and with an empirical basis, we reduce heterogeneity to two groups; within each group, agents are identical. ${ }^{6}$ We seek to identify the conditions under which there exists a linear equilibrium with symmetric treatment of agents in the same group (i.e., we are looking for equilibria such that demand functions are both linear and identical among individuals of the same type). After showing that any such equilibrium must be unique, we derive comparative statics results.

More specifically, our analysis establishes that the number of group members, the transactions costs, the extent to which bidders' valuations are correlated, and the precision of private information affect the sensitivity of traders' demands to private information and prices. When valuations are more correlated, traders react less to the private signal and to the price. We also find that the relative market power of a group increases with the precision of its private information and decreases with its transaction costs. Furthermore, if the transaction costs of a group increase, then the traders of the other group respond strategically by diminishing their reaction to private information and submitting steeper schedules. This result is consistent with the empirical findings of Cassola et al. (2013) in European post-crisis liquidity auctions.

If a group of traders is stronger in the sense described previously (i.e., if its private information is more precise, its transaction costs are lower, and it is more oligopolistic), then the members of that group react more (than do the bidders of the other group) to the private signal and also to the price. This result may help explain the finding of Hortaçsu and Puller (2008) for the Texas balancing market where, there is no accounting for private information on costs that, small firms use steeper schedules than the theory would predict. ${ }^{7}$

We also find when the expected valuations between groups differ that the auction's expected revenue needs not be decreasing in the transaction costs of bidders, the noise in their signals, or the correlation of values. These findings contrast with the results obtained when groups are symmetric. We bound the expected revenue of the auction between the revenues of auctions involving extremal yet symmetric groups.

In this paper we consider large markets and find that, if there is both a small and a large group of bidders, then the former (oligopsonistic) group has more market power and yet even

Kastl 2012; for a a theoretical model see Boyarchenko et al. 2015).
${ }^{6}$ All traders in a group receive the same signal. In case one group does not receive a signal we reproduce the information structure in Grossman and Stiglitz (1980) of uninformed and informed traders.
${ }^{7}$ Linear supply function models have been used extensively for estimating market power in wholesale electricity auctions. Holmberg et al. (2013) provide a foundation for the continuous approach as an approximation to the discrete supply bids in a spot market. In their experimental work, Brandts et al. (2014) find that observed behavior is more consistent with a supply function model than with a discrete multi-unit auction model. Ciarreta and Espinosa (2010) use Spanish data in finding more empirical support for the smooth supply model than the discrete-bid auction model.
the latter (large) group does not behave competitively since it retains some market power due to incomplete information. We also prove that the equilibrium under imperfect competition converges to a price-taking equilibrium in the limit as the number of traders (of both groups) becomes large.

Finally, we provide a welfare analysis. Toward that end, we characterize the deadweight loss at the equilibrium and show how a subsidy scheme may induce an efficient allocation. We find that if one group is stronger (as previously defined), then it should garner a higher per capita subsidy rate; the reason is that traders in the stronger group will behave more strategically and so must be compensated more to become competitive. The paper also underscores how the bidder heterogeneity in terms of information, preferences, or group size documented in previous work may increase deadweight losses. In particular, when the strong group values the asset at least as much as the weak group, the deadweight loss increases with the quantity auctioned and also with the degree of payoff asymmetries.

Our work is related to the literature on divisible good auctions. Results in symmetric pure common value models have been obtained by Wilson (1979), Back and Zender (1993), and Wang and Zender (2002), among others. ${ }^{8}$

Results in interdependent values models with symmetric bidders are obtained by Vives (2011, 2014) and Ausubel et al. (2014), for example. ${ }^{9}$ Vives (2011), while focusing on the tractable family of linear-Gaussian models, shows how increased correlation in traders' valuations increases the market power of those traders. Bergemann et al. (2015) generalize the information structure in Vives (2011) while retaining the assumption of symmetry. Rostek and Weretka (2012) partially relax that assumption and replace it with a weaker "equicommonality" assumption on the matrix correlation among the agents' values. ${ }^{10} \mathrm{Du}$ and Zhu (2015) consider a dynamic

[^3]auction model with ex post equilibria. For the case of complete information, progress has been made in divisible good auction models by characterizing linear supply function equilibria (e.g., Klemperer and Meyer 1989; Akgün 2004; Anderson and Hu 2008). An exception that incorporates incomplete information is the paper by Kyle (1989), who considers a Gaussian model of a divisible good double auction in which some bidders are privately informed and others are uninformed. Sadzik and Andreyanov (2016) study the design of robust exchange mechanisms in a two-type model similar to the one we present here.

Despite the importance of bidder asymmetry, results in multi-unit auctions have been difficult to obtain. As a consequence, most papers that deal with this issue focus on auctions for a single item. In sealed-bid, first-price, single-unit auctions, an equilibrium exists under quite general conditions (Lebrun 1996; Maskin and Riley 2000a; Athey 2001; Reny and Zamir 2004). Uniqueness is explored in Lebrun (1999) and Maskin and Riley (2003). Maskin and Riley (2000b) study asymmetric auctions, and Cantillon (2008) shows that the seller's expected revenue declines as bidders become less symmetric. On the multi-unit auction front, progress in establishing the existence of monotone equilibria has been made by McAdams (2003, 2006); those papers address uniform-price auctions characterized by multi-unit demand, interdependent values and independent types. ${ }^{11}$ Reny (2011) stipulates more general existence conditions that allow for infinite-dimensional type and action spaces; these conditions apply to uniformprice, multi-unit auctions with weakly risk-averse bidders and interdependent values (and where bids are restricted to a finite grid).

The rest of our paper is organized as follows. Section 2 outlines the model. Section 3 characterizes the equilibrium, analyzes its existence and uniqueness, and derives comparative statics results. We address large markets in Section 4 and develop the welfare analysis in Section 5. Section 6 concludes. Proofs are gathered in the Appendix.

## 2 The model

Traders, of whom there are a finite number, face an inelastic supply for a risky asset. Let $Q$ denote the aggregate quantity supplied in the market. In this market there are buyers of two types: type 1 and type 2 . Suppose that there are $n_{i}$ traders of type $i, i=1,2$. In that case, if the asset's price is $p$, then the profits of a representative type- $i$ trader who buys $x_{i}$ units of the of private signals. The equilibrium they derive is therefore still symmetric because all traders use identical strategies.
${ }^{11}$ McAdams (2006) uses a discrete bid space and atomless types to show that, with risk neutral bidders, monotone equilibria exist. The demonstration is based on checking that the single-crossing condition used by Athey (2001) for the single-object case extends to multi-unit auctions.
asset are given by

$$
\pi_{i}=\left(\theta_{i}-p\right) x_{i}-\lambda_{i} x_{i}^{2} / 2
$$

So, for any trader of type $i$, the marginal benefit of buying $x_{i}$ units of the asset is $\theta_{i}-\lambda_{i} x_{i}$, where $\theta_{i}$ denotes the valuation of the asset and $\lambda_{i}>0$ reflects an adjustment for transaction costs or opportunity costs (or a proxy for risk aversion). Traders maximize expected profits and submit demand schedules, after which the auctioneer selects a price that clears the market. The case of supply schedule competition for inelastic demand is easily accommodated by considering negative demands ( $x_{i}<0$ ) and a negative inelastic supply ( $Q<0$ ). In this case, a producer of type $i$ has a quadratic production cost $-\theta_{i} x_{i}+\lambda_{i} x_{i}^{2} / 2$.

We assume that $\theta_{i}$ is normally distributed with mean $\bar{\theta}_{i}$ and variance $\sigma_{\theta}^{2}, i=1,2$. The random variables $\theta_{1}$ and $\theta_{2}$ may be correlated, with correlation coefficient $\rho \in[0,1]$. Therefore, $\operatorname{cov}\left(\theta_{1}, \theta_{2}\right)=\rho \sigma_{\theta}^{2} .{ }^{12}$ All type- $i$ traders receive the same noisy signal $s_{i}=\theta_{i}+\varepsilon_{i}$, where $\varepsilon_{i}$ is normally distributed with null mean and variance $\sigma_{\varepsilon_{i}}^{2}$. Error terms in the signals are uncorrelated across groups $\left(\operatorname{cov}\left(\varepsilon_{1}, \varepsilon_{2}\right)=0\right)$ and are also uncorrelated with valuations of the asset $\left(\operatorname{cov}\left(\varepsilon_{i}, \theta_{j}\right)=0, i, j=1,2\right)$.

In our model, two traders of distinct types may differ in several respects:

- different willingness to possess the asset $\left(\theta_{1} \neq \theta_{2}\right)$,
- different transaction costs $\left(\lambda_{1} \neq \lambda_{2}\right)$, and/or
- different levels of precision of private information $\left(\sigma_{\varepsilon_{1}}^{2} \neq \sigma_{\varepsilon_{2}}^{2}\right)$.

Applications of this model are Treasury auctions and liquidity auctions. For Treasury auctions, $\theta_{i}$ is the private value of the securities to a bidder of type $i$; that value incorporates not only the resale value but also idiosyncratic elements as different liquidity needs of bidders in the two groups. For liquidity auctions, $\theta_{i}$ is the price (or interest rate) that group $i$ commands in the secondary interbank market (which is over-the-counter). Here $\lambda_{i}$ reflects the structure of a counterparty's pool of collateral in a repo auction. A bidder bank prefers to offer illiquid collateral to the central bank in exchange for funds; as allotments increase, however, the bidder must offer more liquid types of collateral which have a higher opportunity cost.

## 3 Equilibrium

Denote by $X_{i}$ the strategy of a type- $i$ bidder, $i=1,2$, which is a mapping from the signal space to the space of demand functions. Thus, $X_{i}\left(s_{i}, \cdot\right)$ is the demand function of a type- $i$ bidder

[^4]that corresponds to a given signal $s_{i}$. Given her signal $s_{i}$, each bidder in a Bayesian equilibrium chooses a demand function that maximizes her conditional profit (while taking as given the other traders' strategies). Our attention will be restricted to anonymous linear Bayesian equilibria in which strategies are linear and identical among traders of the same type (for short, equilibria).

Definition. An equilibrium is a linear Bayesian equilibrium such that the demand functions for traders of type $i, i=1,2$, are identical and equal to

$$
X_{i}\left(s_{i}, p\right)=b_{i}+a_{i} s_{i}-c_{i} p
$$

where $b_{i}, a_{i}$, and $c_{i}$ are constants.

### 3.1 Equilibrium characterization

Consider a trader of type $i$. If rival's strategies are linear and identical among traders of the same type and if the market clears, that is, if $\left(n_{i}-1\right) X_{i}\left(s_{i}, p\right)+x_{i}+n_{j} X_{j}\left(s_{j}, p\right)=Q$, for $j=1,2$ and $j \neq i$, then this trader faces the residual inverse supply $p=I_{i}+d_{i} x_{i}$, where $I_{i}=\left(\left(n_{i}-1\right)\left(b_{i}+a_{i} s_{i}\right)+n_{j}\left(b_{j}+a_{j} s_{j}\right)-Q\right) /\left(\left(n_{i}-1\right) c_{i}+n_{j} c_{j}\right)$ and $d_{i}=1 /\left(\left(n_{i}-1\right) c_{i}+n_{j} c_{j}\right)$. The slope $\left(d_{i}\right)$ is an index of the trader's market power. ${ }^{13}$ As a consequence, this trader's information set $\left(s_{i}, p\right)$ is informationally equivalent to $\left(s_{i}, I_{i}\right)$. The bidder therefore chooses $x_{i}$ to maximize

$$
\mathbb{E}\left[\pi_{i} \mid s_{i}, p\right]=\left(\mathbb{E}\left[\theta_{i} \mid s_{i}, I_{i}\right]-I_{i}-d_{i} x_{i}\right) x_{i}-\lambda_{i} x_{i}^{2} / 2 .
$$

The first-order condition (FOC) is given by $\mathbb{E}\left[\theta_{i} \mid s_{i}, I_{i}\right]-I_{i}-2 d_{i} x_{i}-\lambda_{i} x_{i}=0$, or equivalently,

$$
\begin{equation*}
X_{i}\left(s_{i}, p\right)=\left(\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p\right) /\left(d_{i}+\lambda_{i}\right) \tag{1}
\end{equation*}
$$

The second-order condition (SOC) that guarantees a maximum is $2 d_{i}+\lambda_{i}>0$, which implies that $d_{i}+\lambda_{i}>0$. Using the expression for $I_{i}$ and assuming that $a_{j} \neq 0$, we can show that ( $s_{i}, p$ ) is informationally equivalent to $\left(s_{1}, s_{2}\right)$. Therefore, since $\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]=\mathbb{E}\left[\theta_{i} \mid s_{i}, I_{i}\right]$, it follows that

$$
\begin{equation*}
\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]=\mathbb{E}\left[\theta_{i} \mid s_{1}, s_{2}\right] . \tag{2}
\end{equation*}
$$

According to Gaussian distribution theory,

$$
\begin{equation*}
\mathbb{E}\left[\theta_{i} \mid s_{i}, s_{j}\right]=\bar{\theta}_{i}+\Xi_{i}\left(s_{i}-\bar{\theta}_{i}\right)+\Psi_{i}\left(s_{j}-\bar{\theta}_{j}\right) \tag{3}
\end{equation*}
$$

where

$$
\Xi_{i}=\frac{1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{j}}^{2}}{\left(1+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{j}}^{2}\right)-\rho^{2}} \text { and } \Psi_{i}=\frac{\rho \widehat{\sigma}_{\varepsilon_{i}}^{2}}{\left(1+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{j}}^{2}\right)-\rho^{2}}
$$

with $\widehat{\sigma}_{\varepsilon_{i}}^{2}=\sigma_{\varepsilon_{i}}^{2} / \sigma_{\theta}^{2}$ and $\widehat{\sigma}_{\varepsilon_{j}}^{2}=\sigma_{\varepsilon_{j}}^{2} / \sigma_{\theta}^{2}$. We remark that Equation (3) has the following implications.

[^5]1. The private signal $s_{i}$ is useful for predicting $\theta_{i}$ whenever $1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{j}}^{2} \neq 0$, that is, when either the liquidation values are not perfectly correlated $(\rho \neq 1)$ or type- $j$ traders are imperfectly informed about $\theta_{j}\left(\sigma_{\varepsilon_{j}}^{2} \neq 0\right)$.
2. The private signal $s_{j}$ is useful for predicting $\theta_{i}$ whenever $\rho \widehat{\sigma}_{\varepsilon_{i}}^{2} \neq 0$, that is, when the private liquidation values are correlated $(\rho \neq 0)$ and type- $i$ traders are imperfectly informed about $\theta_{i}\left(\sigma_{\varepsilon_{i}}^{2} \neq 0\right)$.

Our first proposition summarizes the previous results. It shows the relationship between $a_{i}$ and $c_{i}$ in equilibrium and also indicates that these coefficients are positive.

Proposition 1. Let $\rho<1$. In equilibrium, the demand function of a trader of type $i$ $(i=1,2), X_{i}\left(s_{i}, p\right)=b_{i}+a_{i} s_{i}-c_{i} p$, is given by $X_{i}\left(s_{i}, p\right)=\left(\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p\right) /\left(d_{i}+\lambda_{i}\right)$, with $d_{i}+\lambda_{i}>0, d_{i}=1 /\left(\left(n_{i}-1\right) c_{i}+n_{j} c_{j}\right)$, and $a_{i}=\Delta_{i} c_{i}>0$, for $\Delta_{i}=1 /\left(1+(1+\rho)^{-1} \widehat{\sigma}_{\varepsilon_{i}}^{2}\right)$. The coefficient $c_{i}$ can be expressed as a function of the ratio $z=c_{1} / c_{2}$, and $z$ is the unique positive solution of a cubic polynomial $p(z)=0$.

The equilibrium demand function depends on $\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]$. As for the price coefficient (see Expression (11) in the Appendix), $c_{i}=\left(1-\Psi_{i}\left(n_{i} c_{i}+n_{j} c_{j}\right)\left(n_{j} a_{j}\right)^{-1}\right) /\left(d_{i}+\lambda_{i}\right)$, note that the term $\Psi_{i}\left(n_{i} c_{i}+n_{j} c_{j}\right)\left(n_{j} a_{j}\right)^{-1}$ is the information-sensitivity weight of the price. Note also that, the more informative the price (higher $\left.\Psi_{i}\left(n_{i} c_{i}+n_{j} c_{j}\right)\left(n_{j} a_{j}\right)^{-1}\right)$, the lower the price coefficient (lower $c_{i}$ ). Furthermore, this term vanishes when $\Psi_{i}=0$, that is, when either the valuations are uncorrelated ( $\rho=0$ ) or the private signal $s_{i}$ is perfectly informative $\left(\sigma_{\varepsilon_{i}}^{2}=0\right)$ since in those cases the price conveys no additional information to a trader of type $i$.

Since $a_{i}>0$ and $c_{i}>0$, for $i=1,2$, it follows that in equilibrium the higher the value of the trader's observed private signal (or the lower the price), the higher the quantity she will demand. When $\sigma_{\varepsilon_{i}}^{2}>0$ we have $a_{i}<c_{i}$, since $\Delta_{i}<1$ in this case; when $\sigma_{\varepsilon_{i}}^{2}=0$, we have $\Delta_{i}=1$ and $a_{i}=c_{i}$. Observe that we can write the demand as $X_{i}\left(s_{i}, p\right)=b_{i}+c_{i}\left(\Delta_{i} s_{i}-p\right)$.

Because $p$ is a linear function of $s_{1}$ and $s_{2}$, for $i=1,2$ we have $\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]=\mathbb{E}\left[\theta_{i} \mid s_{1}, s_{2}\right]$ (i.e., Equation (2) holds). The equilibrium price is therefore privately revealing, in other words, the private signal and the price enable a type- $i$ trader to learn as much as about $\theta_{i}$ if she had access to all the information available in the market, $\left(s_{1}, s_{2}\right)$.

If $\rho=0$ or if both signals are perfectly informative $\left(\sigma_{\varepsilon_{i}}^{2}=0, i=1,2\right)$, then bidders do not learn about $\theta_{i}$ from prices. Hence, $\mathbb{E}\left[\theta_{i} \mid s_{i}\right]=\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]=\mathbb{E}\left[\theta_{i} \mid s_{1}, s_{2}\right]$ for $i=1,2$. The demand functions are given by

$$
X_{i}\left(s_{i}, p\right)=\left(\mathbb{E}\left[\theta_{i} \mid s_{i}\right]-p\right) /\left(d_{i}+\lambda_{i}\right), i=1,2 .
$$

Hence, $c_{i}=1 /\left(d_{i}+\lambda_{i}\right)$ and so, given our expression for $d_{i}$, we have

$$
d_{i}=1 /\left(\left(n_{i}-1\right) /\left(d_{i}+\lambda_{i}\right)+n_{j} /\left(d_{j}+\lambda_{j}\right)\right) \text { for } i, j=1,2 \text { and } j \neq i
$$

We can show that, this system has a unique solution satisfying the inequality $d_{i}+\lambda_{i}>0, i=1,2$, iff $n_{1}+n_{2} \geq 3$. In this case, the equilibrium coincides with the full-information equilibrium (denoted by superscript $f) .{ }^{14}$ Furthermore, when $\rho=0$, market power $d_{i}$ is independent of $\sigma_{\varepsilon_{i}}^{2}$, $i=1,2$; and when $\sigma_{\varepsilon_{i}}^{2}=0$, for $i=1,2$, market power is independent of $\rho$.

Our next proposition describes the condition under which an equilibrium exists and shows that, if an equilibrium does exist, then it is unique.

Proposition 2. There exists a unique equilibrium if and only if $\bar{z}_{N}>\bar{z}_{D}$, where $\bar{z}_{N}$ and $\bar{z}_{D}$ denote the highest root of, respectively, $q_{N}(z)$ and $q_{D}(z)$, with

$$
\begin{aligned}
& q_{N}(z)=n_{2}^{2} \Xi_{1} \Delta_{1}^{-1}+n_{2}\left(\Xi_{1} \Delta_{1}^{-1}\left(2 n_{1}-1\right)-\left(n_{1}+1\right)\right) z-\left(n_{1}-1\right)\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} z^{2} \text { and } \\
& q_{D}(z)=-n_{2}\left(n_{2}-1\right)\left(1-\Xi_{2} \Delta_{2}^{-1}\right)+n_{1}\left(\Xi_{2} \Delta_{2}^{-1}\left(2 n_{2}-1\right)-\left(n_{2}+1\right)\right) z+n_{1}^{2} \Xi_{2} \Delta_{2}^{-1} z^{2} .
\end{aligned}
$$

Let $z=c_{1} / c_{2}$. Then, in equilibrium, $\bar{z}_{D}<z<\bar{z}_{N}, \lim _{\lambda_{1} \rightarrow 0} z=\bar{z}_{N}$ and $\lim _{\lambda_{2} \rightarrow 0} z=\bar{z}_{D}$.
For an equilibrium to exist we must have $c_{i}>0(i=1,2)$ and these inequalities hold if and only if (iff) $\bar{z}_{D}<z<\bar{z}_{N}$. No equilibrium exists for $\rho$ close to 1 or for low $n_{1}+n_{2}$. Neither does an equilibrium exist when $\rho=1$. If the price reveals a sufficient statistic for the common valuation, then no trader has an incentive to place any weight on her signal. But if traders put no weight on signals, then the price contains no information about the common valuation. This conundrum is related to the Grossman-Stiglitz (1980) paradox.

Remark 1. If $n_{1}=1$ and $n_{2}=1$, then $\bar{z}_{N}=1 /\left(2 \Delta_{1} \Xi_{1}^{-1}-1\right)$ and $\bar{z}_{D}=2 \Delta_{2} \Xi_{2}^{-1}-1$. Since $\Delta_{i} \Xi_{i}^{-1}>1, i=1,2$, we can use direct computation to obtain $\bar{z}_{N}<\bar{z}_{D}$. Applying Proposition 2 , we conclude that no equilibrium exists in this case. Therefore, the inequality $n_{1}+n_{2} \geq 3$ is a necessary condition for the existence of an equilibrium in our model. This result is in line with Kyle (1989) and Vives (2011). ${ }^{15}$

To develop a better understanding of the equilibrium and the condition that guarantees its existence, we consider two particular cases of the model: a monopsony competing with a fringe; and symmetric groups.

[^6]
## Monopsony competing with fringe

Corollary 1. For $n_{2}=1$ the equilibrium exists if $1-\rho^{2}>(2 \rho-1) \widehat{\sigma}_{\varepsilon_{1}}^{2}$ and $n_{1}>$ $\bar{n}_{1}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)$, where $\bar{n}_{1}$ increases with $\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}$, and $\widehat{\sigma}_{\varepsilon_{2}}^{2}$.If, also, $\lambda_{2}=0$ and $\sigma_{\varepsilon_{2}}^{2}=0$, then $\bar{n}_{1}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, 0\right)=1+\rho \widehat{\sigma}_{\varepsilon_{1}}^{2} /\left(1-\rho^{2}-(2 \rho-1) \widehat{\sigma}_{\varepsilon_{1}}^{2}\right)$ and $x_{2}=c_{2}\left(\theta_{2}-p\right)$, with $c_{2}=n_{1} c_{1}$.

An equilibrium with linear demand functions exists provided there is a sufficiently competitive trading environment ( $n_{1}$ high enough). In the particular case where $\lambda_{2}=0$ and $\sigma_{\varepsilon_{2}}^{2}=0$, expressions for the equilibrium coefficients can be characterized explicitly (see the Appendix). From the expressions for $c_{i}(i=1,2)$ it follows that, if $n_{1}=\bar{n}_{1}$, then the equilibrium cannot exist because in this case the demand functions would be completely inelastic ( $c_{i}=0, i=1,2$ ).

## Symmetric groups

Consider the following symmetric case: $n_{2}=n_{1}=n, \lambda_{1}=\lambda_{2}=\lambda$, and $\sigma_{\varepsilon_{1}}^{2}=\sigma_{\varepsilon_{2}}^{2}=\sigma_{\varepsilon}^{2}$. Here $z=1$ in equilibrium. From Proposition 2 we know that, if an equilibrium exists, then the value of $z$ is in the interval $\left(\bar{z}_{D}, \bar{z}_{N}\right)$. It follows that $\bar{z}_{N}>1>\bar{z}_{D}$ or, equivalently, that $q_{N}(1)>0$ and $q_{D}(1)>0$. After performing some algebra, we find that the foregoing inequalities are satisfied iff $n>1+\rho \widehat{\sigma}_{\varepsilon}^{2} /\left((1-\rho)\left(1+\rho+\widehat{\sigma}_{\varepsilon}^{2}\right)\right)$, where $\widehat{\sigma}_{\varepsilon}^{2}=\sigma_{\varepsilon}^{2} / \sigma_{\theta}^{2}$. Therefore, the equilibrium's existence is guaranteed provided either that $n$ is high enough or that $\rho$ or $\widehat{\sigma}_{\varepsilon}^{2}$ is low enough.

Vives (2011) also analyzes divisible good auctions with symmetric bidders, but in his model the bidders receive different private signals. The condition that guarantees existence of an equilibrium in Vives' setup is $2 n>2+M$, where $M=2 n \rho \widehat{\sigma}_{\varepsilon}^{2} /\left((1-\rho)\left(1+(2 n-1) \rho+\widehat{\sigma}_{\varepsilon}^{2}\right)\right)$. Direct computation yields that the condition derived in the model of Vives is more stringent than the condition derived in our setup. The reason is that, in Vives (2011), the degree of asymmetry in information (and induced market power) is greater because each of the $2 n$ traders receives a private signal.

The rest of this subsection is devoted to describing some properties that satisfy the equilibrium coefficients and then to comparing the equilibrium quantities.

## Informed and uninformed groups

Consider the case where group 1 is uninformed (with no signal) while group 2 is informed $\left(\sigma_{\varepsilon_{2}}^{2}<\infty\right)$. This is akin to the information structure considered by Grossman and Stiglitz (1980). Then equilibrium exists if and only if $\rho<\frac{n_{2}\left(n_{1}+n_{2}-2\right)}{2 n_{1}+n_{2}-2}$. We have that $a_{2}=\left(1+\frac{\sigma_{\varepsilon_{2}}^{2}}{\sigma_{\theta}^{2}}\right)^{-1} c_{2}$. This expression is different from the case in which the group 1 is informed since then $a_{2}=$ $\left(1+\frac{\sigma_{\varepsilon_{2}}^{2}}{(1+\rho) \sigma_{\theta}^{2}}\right)^{-1} c_{2}$. Therefore, the equilibrium that we obtain here is not obtained as a limit of the general case when $\sigma_{\varepsilon_{1}}^{2}$ converges to infinity, except in the case in which $\rho=0$. This is so,
since when $\rho>0$, no matter how large $\sigma_{\varepsilon_{1}}^{2}$ is, but finite, there is learning from the price by the two groups. By contrast, when group 1 does not receive any private signal, there is no learning from the price by group 2. Furthermore, in this last case, we have that $c_{1}$ and $c_{2}\left(d_{1}\right.$ and $\left.d_{2}\right)$ are independent of $\sigma_{\varepsilon_{2}}^{2}$, and $b_{2}$ is independent of $\bar{\theta}_{1}$ and $Q$.

In the case where group receives a noisy signal $\left(\sigma_{\varepsilon_{1}}^{2}<\infty\right)$ and group 2 is perfectly informed $\left(\sigma_{\varepsilon_{2}}^{2}=0\right)$, then $b_{2}=0$ and $a_{1}=c_{1}$.

## Comparative statics

We start by considering how the model's underlying parameters affect the equilibrium and, in particular, market power (Proposition 3). We then explore how the equilibrium is affected when there are two distinct groups of traders, that is, a strong group and a weak group (Corollary 2).

Proposition 3. Let $\rho \sigma_{\varepsilon_{1}}^{2} \sigma_{\varepsilon_{2}}^{2}>0$. Then, for $i=1,2, i \neq j$, the following statements hold.
(i) An increase in $\bar{\theta}_{i}$ or $Q$, or a decrease in $\bar{\theta}_{j}$, raises the demand intercept $b_{i}$.
(ii) An increase in $\lambda_{i}, \lambda_{j}, \sigma_{\varepsilon_{i}}^{2}, \sigma_{\varepsilon_{j}}^{2}$, or $\rho$ makes demand less responsive to private signals and prices (lower $a_{i}$ and $c_{i}$ ) and increases market power $\left(d_{i}\right)$.
(iii) If $\sigma_{\varepsilon_{i}}^{2}$ and/or $\lambda_{i}$ increase, then $d_{i} / d_{j}$ decreases.
(iv) If $n_{i}$ and/or $n_{j}$ increase, then $d_{i}$ decreases.

Remark 2. If $\rho=0$, then: (a) both $c_{i}$ and $d_{i}$ (as well as $c_{j}$ and $a_{j}, j \neq i$ ) are independent of $\sigma_{\varepsilon_{i}}^{2}$; (b) $a_{i}$ decreases with $\sigma_{\varepsilon_{i}}^{2}$; and (c) $b_{i}$ is independent of both $Q$ and $\bar{\theta}_{j}$. If $\sigma_{\varepsilon_{i}}^{2}=0$ for $i=1,2$, then $b_{i}=0$, and $c_{i}, c_{j}, a_{i}, a_{j}, d_{i}$, and $d_{j}, i=1,2, j \neq i$, are independent of $\rho$. That is, for the information parameters to matter for market power, it is necessary that prices convey information. And given that the equilibrium values of $d_{1}$ and $d_{2}$ when $\rho=0$ (or when $\sigma_{\varepsilon_{i}}^{2}=0$, $i=1,2$ ) are equal to those corresponding to the full-information setup, Proposition 3(ii) implies that, if $\rho \sigma_{\varepsilon_{1}}^{2} \sigma_{\varepsilon_{2}}^{2}>0$, then $d_{i}^{f}<d_{i}, i=1,2$. Thus, in this case asymmetric information increases the market power of traders in both groups beyond the full-information level.

By Lemma A1, the only equilibrium coefficient affected by the quantity offered in the auction $(Q)$ and by the prior mean of the valuations $\left(\bar{\theta}_{i}\right.$ and $\left.\bar{\theta}_{j}\right)$ is the coefficient $b_{i}$. Proposition 3(i) indicates that if $Q$ increases, then all the bidders will increase their demand (higher $b_{1}$ and $b_{2}$ ). Moreover, if the prior mean of the valuation of group $i$ increases, then the bidders in this group will demand a greater quantity of the risky asset (higher $b_{i}$ ). Then the intercept of the inverse residual supply for the group $j$ bidder rises in response to a higher $\bar{\theta}_{i}$. That reaction leads the traders in group $j$ to reduce their demand for the risky asset (lower $b_{j}$ ).

Part (ii) of Proposition 3 shows how the response to private information and price varies with several parameters. If the transaction costs for a bidder increase, then that bidder is less interested in the risky asset and so $a_{i}$ and $c_{i}$ are each decreasing in $\lambda_{i}$. Moreover, any increase in
a group's transaction costs also affects the behavior of traders in the other group. If $\lambda_{i}$ increases, then $c_{i}$ decreases, in which case the slope of the inverse residual supply for group $j$ increases (higher $d_{j}$ ). This change leads group- $j$ traders to reduce their demand sensitivity to signals and prices (lower $a_{j}$ and $c_{j}$ ). We can therefore see how an increase in the transaction costs for group- $i$ traders (say, a deterioration of their collateral in liquidity auctions that raises $\lambda_{i}$ ) leads not only to steeper demands for bidders in group $i$ but also, as a reaction, to steeper demands for group- $j$ traders. Figure 1 illustrates the case of initially identical groups that become differentiated after a shock induces a higher $\lambda_{1}$ and also raises group's willingness to pay for liquidity as in a crisis situation (both $\bar{\theta}_{1}$ and $\bar{\theta}_{2}$, which affect the intercepts of the demand functions).


Figure 1: Equilibrium demand functions for $\rho=0.75, \sigma_{\theta}^{2}=5, Q=4, n_{i}=5, \sigma_{\varepsilon_{i}}^{2}=1$, and $s_{i}=\bar{\theta}_{i}, i=1,2$.

We also analyze how the response to private information and price varies with a change in the precision of private signals. If the private signal of type- $i$ bidders is less precise (higher $\sigma_{\varepsilon_{i}}^{2}$ ), then their demand is less sensitive to private information and prices. Thus a trader finds it optimal to rely less on her private information when her private signal is less precise. A private signal of reduced precision also gives the type- $i$ bidder more incentive to consider prices when predicting $\theta_{i}$, which leads in turn to this bidder having a steeper demand function (lower $c_{i}$ ). The same can be said for a bidder of type $j$ because of strategic complementarity in the slopes of demand functions. ${ }^{16}$

[^7]We also find that the more highly the valuations are correlated (higher $\rho$ ), the less is trader responsiveness to private signals (lower $a_{i}, i=1,2$ ) and the steeper are inverse demand functions (lower $c_{i}, i=1,2$ ). We can explain these results by recalling that, when the valuations are correlated $(\rho>0)$, a type- $i$ trader learns about $\theta_{i}$ from prices. In fact, the price is more informative about $\theta_{i}$ when $\rho$ is larger, in which case demand is less sensitive to private information. The rationale for the relationship between the correlation coefficient ( $\rho$ ) and the slopes of demand functions is as follows. An increment in the price of the risky asset makes an agent more optimistic about her valuation, which leads to less of a reduction in demand quantity than in the case of uncorrelated valuations. ${ }^{17}$

Proposition 3(iii) states that any increase in the signal's noise or in group $i$ 's transaction costs has the effect of reducing its relative market power, since then the ratio $d_{i} / d_{j}(i \neq j)$ decreases. Finally, part (iv) formalizes the anticipated result that an increase in the number of auction participants (higher $n_{i}$ or $n_{j}$ ) reduces the market power of traders in both groups.

Corollary 2. Suppose that $\sigma_{\varepsilon_{1}}^{2} \geq \sigma_{\varepsilon_{2}}^{2}, \lambda_{1} \geq \lambda_{2}$, and $n_{1} \geq n_{2}$, and suppose that at least one of these inequalities is strict. Then, in equilibrium, the following statements hold.
(i) The stronger group (here, group 2) reacts more both to private information and to prices ( $a_{1}<a_{2}, c_{1}<c_{2}$ ) and has more market power $\left(d_{1}<d_{2}\right)$ than does the weaker group.
(ii) The value of the difference $d_{1}+\lambda_{1}-\left(d_{2}+\lambda_{2}\right)$ is, in general, ambiguous. If

$$
\begin{equation*}
\frac{(1-\rho) n_{1} n_{2}\left(1+\rho+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)}{n_{2}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)+n_{1} \rho \widehat{\sigma}_{\varepsilon_{1}}^{2}}+\frac{(1-\rho) n_{1}\left(n_{2}-1\right)\left(1+\rho+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)}{n_{1}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)+n_{2} \rho \widehat{\sigma}_{\varepsilon_{2}}^{2}} \leq 1 \tag{4}
\end{equation*}
$$

then $d_{1}+\lambda_{1}<d_{2}+\lambda_{2}$ always holds. Otherwise, $d_{1}+\lambda_{1}>d_{2}+\lambda_{2}$ iff $\lambda_{1} / \lambda_{2}$ is high enough.
Part (i) of this corollary shows that if a group of traders is less informed, has higher transaction costs, and is more numerous, then it reacts less both to private signals and to prices. Observe in particular that group-1 traders, having less precise private information, rely more on the price for information (higher $\left.\Psi_{1}\left(n_{1} c_{1}+n_{2} c_{2}\right)\left(n_{2} a_{2}\right)^{-1}\right)$; as a result, their overall price response $\left(c_{1}=\left(1-\Psi_{1}\left(n_{1} c_{1}+n_{2} c_{2}\right)\left(n_{2} a_{2}\right)^{-1}\right) /\left(d_{1}+\lambda_{1}\right)\right)$ is smaller. Similarly, group- 1 traders, for whom $n_{1}$ is larger, put more information weight on the price (which depends more strongly on $s_{1}$ ).

Corollary 2(ii) is useful for comparing allocations across groups. It indicates that the inequality $d_{1}+\lambda_{1}>d_{2}+\lambda_{2}$ holds whenever (a) the differences between groups stem mainly from transaction costs; ${ }^{18}$ and (b) $\lambda_{1} / \lambda_{2}$ is high enough. If signals are perfect $\left(\sigma_{\varepsilon_{i}}^{2}=0, i=1,2\right)$ or if

[^8]$\rho=0$, then part (i) holds and $d_{1}+\lambda_{1}>d_{2}+\lambda_{2}$ iff $\lambda_{1}>\lambda_{2}$.

## Equilibrium quantities

Finally, we examine the equilibrium quantities. Let $t_{i}=\mathbb{E}\left[\theta_{i} \mid s_{1}, s_{2}\right], i=1,2$, be the predicted values with full information $\left(s_{1}, s_{2}\right)$. After some algebra, it follows that equilibrium quantities are functions of the vector of predicted values $t=\left(t_{1}, t_{2}\right)$ :

$$
\begin{equation*}
x_{i}(t)=\underbrace{\frac{n_{j}\left(t_{i}-t_{j}\right)}{n_{i}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(d_{i}+\lambda_{i}\right)}}_{x_{i}^{I}(t)}+\underbrace{\frac{d_{j}+\lambda_{j}}{n_{i}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(d_{i}+\lambda_{i}\right)}}_{x_{i}^{C}(t)} Q, i=1,2, j \neq i . \tag{5}
\end{equation*}
$$

Observe that, according to these expressions, the equilibrium quantities can be decomposed into two terms: a valuation trading term and a clearing trading term, which we denote by (respectively) $x_{i}^{I}(t)$ and $x_{i}^{C}(t)$ for group $i, i=1,2$. With regard to the information trading term, it vanishes when $t_{1}=t_{2}$, but has a positive (resp. negative) value for the group with the higher (resp. lower) value of $t_{i}$. Moreover, $n_{1} x_{1}^{I}(t)+n_{2} x_{2}^{I}(t)=0$. As for the clearing trading term, we remark that it vanishes when $Q=0$; otherwise, it is positive for both groups yet lower (resp. higher) for the group with higher (resp. lower) $d_{i}+\lambda_{i}$. In addition, $n_{1} x_{1}^{C}(t)+n_{2} x_{2}^{C}(t)=Q$.

Taking expectations in Equation (5), we have

$$
\mathbb{E}\left[x_{1}(t)\right]-\mathbb{E}\left[x_{2}(t)\right]=\frac{n_{1}+n_{2}}{n_{1}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(d_{1}+\lambda_{1}\right)}\left(\bar{\theta}_{1}-\bar{\theta}_{2}\right)+\frac{d_{2}+\lambda_{2}-\left(d_{1}+\lambda_{1}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(d_{1}+\lambda_{1}\right)} Q .
$$

Group 1 trades more when it values the asset more highly $\left(\bar{\theta}_{1}>\bar{\theta}_{2}\right)$ and when its traders are less cautious $\left(d_{2}+\lambda_{2}>d_{1}+\lambda_{1}\right)$ than group 2 . By combining Corollary 2 with the equation just displayed, we obtain the following remarks.

Remark 3. If $Q$ is low enough, then $\mathbb{E}\left[x_{1}(t)\right]>\mathbb{E}\left[x_{2}(t)\right]$ whenever $\bar{\theta}_{1}>\bar{\theta}_{2}$. In contrast, if $Q$ is high enough, then $\mathbb{E}\left[x_{1}(t)\right]>\mathbb{E}\left[x_{2}(t)\right]$ whenever $d_{2}+\lambda_{2}>d_{1}+\lambda_{1}$. Under the assumptions of Corollary 2, this latter inequality is satisfied if (4) holds or if $\lambda_{1} / \lambda_{2}$ is sufficiently low.

Remark 4. When $Q=0$ (i.e., the so-called double auction case), then $\mathbb{E}\left[x_{2}(t)\right]<0<$ $\mathbb{E}\left[x_{1}(t)\right]$ iff $\bar{\theta}_{1}>\bar{\theta}_{2}$. Then group 1 consists of buyers and group 2 of sellers.

### 3.2 Bid shading, expected discount, and expected revenue

Our aim here is to identify factors that affect the magnitudes of bid shading, expected discount, and expected revenue. Let $\tilde{t}=\left(n_{1} t_{1}+n_{2} t_{2}\right) /\left(n_{1}+n_{2}\right)$. From the demand of bidders it follows that $p(t)=t_{i}-\left(d_{i}+\lambda_{i}\right) x_{i}(t), i=1,2$. Therefore,

$$
\begin{equation*}
p(t)=\tilde{t}-\left(\left(d_{1}+\lambda_{1}\right) n_{1} x_{1}(t)+\left(d_{2}+\lambda_{2}\right) n_{2} x_{2}(t)\right) /\left(n_{1}+n_{2}\right) . \tag{6}
\end{equation*}
$$

## Bid shading

For a trader of type $i$, the expected marginal benefit of buying $x_{i}$ units of the asset is $t_{i}-\lambda_{i} x_{i}$. Hence, the average marginal benefit is given by $\tilde{t}-\left(\lambda_{1} n_{1} x_{1}+\lambda_{2} n_{2} x_{2}\right) /\left(n_{1}+n_{2}\right)$. The magnitude of bid shading is the difference between the average marginal valuation and the auction price, that is, $\left(d_{1} n_{1} x_{1}+d_{2} n_{2} x_{2}\right) /\left(n_{1}+n_{2}\right)$. We can use Equation (5) to write bid shading as

$$
\begin{equation*}
\frac{n_{2} d_{2}\left(d_{1}+\lambda_{1}\right)+n_{1} d_{1}\left(d_{2}+\lambda_{2}\right)}{\left(n_{1}+n_{2}\right)\left(n_{1}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(d_{1}+\lambda_{1}\right)\right)} Q+\frac{\left(t_{2}-t_{1}\right)\left(d_{2}-d_{1}\right) n_{2} n_{1}}{\left(n_{1}+n_{2}\right)\left(n_{1}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(d_{1}+\lambda_{1}\right)\right)} . \tag{7}
\end{equation*}
$$

At this juncture, some additional remarks are in order.

- Bid shading increases with $Q$.
- When $d_{1}=d_{2}=d$ as in the symmetric case, for instance, bid shading consists of only one term (the first one) and it is equal to $d Q /\left(n_{1}+n_{2}\right)$.
- When $d_{1} \neq d_{2}$, the second term of (7) is negative and bid shading decreases whenever the group that values the asset more highly $\left(t_{i}>t_{j}\right)$ has less market power $\left(d_{i}<d_{j}\right)$.
- If group 1 has higher transaction costs $\left(\lambda_{1}>\lambda_{2}\right)$, is more numerous $\left(n_{1}>n_{2}\right)$, and is less informed $\left(\sigma_{\varepsilon_{1}}^{2}>\sigma_{\varepsilon_{2}}^{2}\right)$ than group 2, then $c_{1}<c_{2}$, and so $d_{1}<d_{2}$. If $t_{1}>t_{2}$, then the second term of (7) is negative and the two terms have opposite signs. Therefore, if $Q$ is low (e.g., $Q=0$ ) or if the difference in predicted values of the asset is high, then negative bid shading obtains.


## Expected discount

The expected discount is defined as $\mathbb{E}[\tilde{t}]-\mathbb{E}[p(t)]$. We can use Equation (6) to write the expected discount as $\left(\left(d_{1}+\lambda_{1}\right) n_{1} \mathbb{E}\left[x_{1}(t)\right]+\left(d_{2}+\lambda_{2}\right) n_{2} \mathbb{E}\left[x_{2}(t)\right]\right) /\left(n_{1}+n_{2}\right)$. Now some algebra yields the following expression for the expected discount:

$$
\begin{equation*}
\frac{\left(d_{1}+\lambda_{1}\right)\left(d_{2}+\lambda_{2}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(d_{1}+\lambda_{1}\right)} Q+\frac{n_{1} n_{2}\left(d_{2}+\lambda_{2}-d_{1}-\lambda_{1}\right)\left(\bar{\theta}_{2}-\bar{\theta}_{1}\right)}{\left(n_{1}+n_{2}\right)\left(n_{1}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(d_{1}+\lambda_{1}\right)\right)} \tag{8}
\end{equation*}
$$

Here our related comments are as follows.

- When $d_{1}+\lambda_{1}=d_{2}+\lambda_{2}=d+\lambda$ (as in the symmetric case), the expected discount is $(d+\lambda) Q /\left(n_{1}+n_{2}\right)$.
- The first term is always positive provided $Q>0$, whereas the second term is positive whenever $\left(d_{2}+\lambda_{2}-d_{1}-\lambda_{1}\right)\left(\bar{\theta}_{2}-\bar{\theta}_{1}\right)>0$. Therefore, the expected discount is lower whenever the group that values the asset more highly $\left(\bar{\theta}_{2}>\bar{\theta}_{1}\right)$ has a lower "total transaction cost" $\left(d_{2}+\lambda_{2}<d_{1}+\lambda_{1}\right)$.
- If group 1 ex ante values the asset more $\left(\bar{\theta}_{1}>\bar{\theta}_{2}\right)$, has higher transaction costs $\left(\lambda_{1}>\lambda_{2}\right)$, is more numerous $\left(n_{1}>n_{2}\right)$, and is less informed $\left(\sigma_{\varepsilon_{1}}^{2}>\sigma_{\varepsilon_{2}}^{2}\right)$, then Corollary 2 shows that $d_{1}+\lambda_{1}>d_{2}+\lambda_{2}$ whenever (a) the differences between groups are due mostly to transaction costs and (b) $\lambda_{1} / \lambda_{2}$ is high enough. In this case, both terms are positive and so the expected discount is positive. Yet, if both groups have similar transactions costs, then the two terms in (8) have opposite signs. In particular, we expect a negative discount when $Q$ is low.


## Expected revenue

The expected price is given by

$$
\mathbb{E}[p]=\left(\frac{n_{1}}{d_{1}+\lambda_{1}} \bar{\theta}_{1}+\frac{n_{2}}{d_{2}+\lambda_{2}} \bar{\theta}_{2}-Q\right) /\left(\frac{n_{1}}{d_{1}+\lambda_{1}}+\frac{n_{2}}{d_{2}+\lambda_{2}}\right) .
$$

It is worth noting that, in the double auction case $(Q=0), \mathbb{E}[p]$ is a convex combination of $\bar{\theta}_{1}$ and $\bar{\theta}_{2}$. Also, for symmetric groups (except possibly with respect to the means) we have $\mathbb{E}[p]=\left(\bar{\theta}_{1}+\bar{\theta}_{2}\right) / 2$.

The seller's expected revenue is $E[p] Q$ and, provided that he has enough supplies at no cost, the revenue-maximizing supply is given by $Q^{*}=\frac{1}{2}\left(\frac{n_{1}}{d_{1}+\lambda_{1}} \bar{\theta}_{1}+\frac{n_{2}}{d_{2}+\lambda_{2}} \bar{\theta}_{2}\right)$. That supply $Q^{*}$ is increasing in the expected valuations and in the number of traders; it is decreasing in those traders' market power and transaction costs.

Proposition 4. Let $\rho \sigma_{\varepsilon_{1}}^{2} \sigma_{\varepsilon_{2}}^{2}>0$. Then, in equilibrium, the following statements hold.
(i) If $\bar{\theta}_{1}=\bar{\theta}_{2}$, then the expected price is increasing in $n_{i}$ but is decreasing in $\lambda_{i}, \sigma_{\varepsilon_{i}}^{2}$, and $\rho$, $i=1,2$. Otherwise, if $\left|\bar{\theta}_{1}-\bar{\theta}_{2}\right|$ is large enough, then these results need not hold.
(ii) The expected revenue:

- increases with $\bar{\theta}_{i}$ for $i=1,2$, and increases with $Q$ for $E[p]>0$;
- is between (a) the larger expected revenue of the auction in which both groups are ex ante identical with a large number of bidders (each group with $\max \left\{n_{1}, n_{2}\right\}$ ), high expected valuation $\left(\max \left\{\bar{\theta}_{1}, \bar{\theta}_{2}\right\}\right)$, low transaction costs $\left(\min \left\{\lambda_{1}, \lambda_{2}\right\}\right)$ and precise signals $\left(\min \left\{\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}\right\}\right)$ and (b) the smaller expected revenue of the auction in which both groups are ex ante identical but with the opposite characteristics (i.e., $\min \left\{n_{1}, n_{2}\right\}$, $\min \left\{\bar{\theta}_{1}, \bar{\theta}_{2}\right\}, \max \left\{\lambda_{1}, \lambda_{2}\right\}$, and $\max \left\{\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}\right\}$ ).

Remark 5. If $\rho=0$, then $\mathbb{E}[p]$ is independent of $\sigma_{\varepsilon_{i}}^{2}(i=1,2)$, and if $\sigma_{\varepsilon_{i}}^{2}=0, i=1,2$, then $\mathbb{E}[p]$ is independent of $\rho$. The reason is that in both cases, $d_{i}$ is independent of $\sigma_{\varepsilon_{i}}^{2}$ and $\rho$. If $\sigma_{\varepsilon_{1}}^{2}=\infty$ then $\mathbb{E}[p]$ is independent of $\sigma_{\varepsilon_{2}}^{2}$.

Proposition 4 indicates that the relationship between expected price (on the one hand) and $\lambda_{i}, \sigma_{\varepsilon_{i}}^{2}$, and $\rho, i=1,2$ (on the other hand) is potentially ambiguous. For example, if $\bar{\theta}_{2}-\bar{\theta}_{1}$ is high enough, then $\mathbb{E}[p]$ is decreasing in $n_{1}$; yet, if $\bar{\theta}_{1}=\bar{\theta}_{2}$, then the derived results are in line with those in the symmetric case, where $\mathbb{E}[p]=\bar{\theta}-(d+\lambda) Q / 2 n$ (see Vives 2010, Prop. 2).

We should like to understand how ex ante differences among bidders affect the seller's expected revenue. Suppose that group 2 is our strong group; it has lower transaction costs $\left(\lambda_{2}<\lambda_{1}\right)$, is less numerous ( $n_{2}<n_{1}$ ), and is better informed ( $\left.\sigma_{\varepsilon_{2}}^{2}<\sigma_{\varepsilon_{1}}^{2}\right)$. If this group values the asset less, $\bar{\theta}_{2}<\bar{\theta}_{1}$ (resp., values it more, $\bar{\theta}_{2}>\bar{\theta}_{1}$ ), then expected revenue is lower (resp., higher) than in the case where $\bar{\theta}_{1}=\bar{\theta}_{2}$. If $\bar{\theta}_{1} \approx \bar{\theta}_{2}$, then Proposition 4(i) suggests that group 2's relatively small size $\left(n_{2}<n_{1}\right)$ reduces the seller's expected revenue, although both its relatively low transaction costs $\left(\lambda_{2}<\lambda_{1}\right)$ and its relatively precise signals ( $\sigma_{\varepsilon_{2}}^{2}<\sigma_{\varepsilon_{1}}^{2}$ ) have the opposite effect. So in general, the ex ante differences between the two groups have an ambiguous effect on the seller's expected revenue. Nonetheless, part (ii) of Proposition 4 directly follows from part (i).

## 4 Large markets

Our objective in this section is to determine whether (or not) the equilibrium under imperfect competition converges to a price-taking equilibrium in the limit as the number of traders becomes large. We examine two possible scenarios: in the first, only group 1 is large; in the second, both groups of bidders are large. The per capita supply (denoted by $q$ ) is assumed to be inelastic, that is, $Q=\left(n_{1}+n_{2}\right) q$.

### 4.1 Oligopsony with competitive fringe

Proposition 5. Let $\rho \sigma_{\varepsilon_{1}}^{2}>0$. Suppose that $n_{1} \rightarrow \infty$ and $n_{2}<\infty$. Then an equilibrium exists iff $n_{2}>\bar{n}_{2}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)$, where $\bar{n}_{2}$ is increasing in $\rho$ and $\widehat{\sigma}_{\varepsilon_{1}}^{2}$ and where $\bar{n}_{2}$ is decreasing in $\widehat{\sigma}_{\varepsilon_{2}}^{2}$ whenever $(2 \rho-1) \widehat{\sigma}_{\varepsilon_{1}}^{2}<1-\rho^{2}$. An agent in the large group absorbs the inelastic per capita supply in the limit $\left(\lim _{n_{1} \rightarrow \infty} b_{1}=q, \lim _{n_{1} \rightarrow \infty} a_{1}=\lim _{n_{1} \rightarrow \infty} c_{1}=0\right)$ and retains some market power $\left(\lim _{n_{1} \rightarrow \infty} d_{1}>0\right)$, while an agent in the small group commands a higher degree of market power $\left(\lim _{n_{1} \rightarrow \infty} d_{2}>\lim _{n_{1} \rightarrow \infty} d_{1}\right)$.

When $n_{2}=1$, the existence condition stated in Proposition 5 boils down to $(2 \rho-1) \widehat{\sigma}_{\varepsilon_{1}}^{2}<$
$1-\rho^{2}$ from Corollary 1. Equation (33) shows that, when $n_{2}=\bar{n}_{2}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)$, the demand functions for bidders in group 2 would be completely inelastic $\left(\lim _{n_{1} \rightarrow \infty} c_{2}=0\right)$. This explains why the inequality $n_{2}>\bar{n}_{2}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)$ is required for the existence of equilibrium. Neither group 1 nor group 2 has flat aggregate demand in the limit, and each group has some market power. We see that an agent in the large group just absorbs the inelastic per capita supply, behaving like a "Cournot quantity setter", and keeping some market power ( $\lim _{n_{1} \rightarrow \infty} d_{1}>0$ ), while bidders in the small group command relatively more market power $\left(\lim _{n_{1} \rightarrow \infty} d_{2}>\lim _{n_{1} \rightarrow \infty} d_{1}\right)$. It is worth to remark that the large group retains market power in the limit only if this group learns from the price (incomplete information and correlation of values, $\rho \sigma_{\varepsilon_{1}}^{2}>0$ ). In this case the aggregate demand of group 1 does not become flat, $\lim _{n_{1} \rightarrow \infty} n_{1} c_{1}<\infty$. If $\rho \sigma_{\varepsilon_{1}}^{2}=0, \lim _{n_{1} \rightarrow \infty} n_{1} c_{1}=\infty$ and there is no market power in the limit: $\lim _{n_{1} \rightarrow \infty} d_{i}=0, i=1,2$. It is easy to see also that, in the limit, the price depends only on the valuations and market power of agents in the competitive fringe: $\lim _{n_{1} \rightarrow \infty} p=\mathbb{E}\left[\theta_{1} \mid s_{1}, s_{2}\right]-\left(\lim _{n_{1} \rightarrow \infty} d_{1}+\lambda_{1}\right) q$ and that the limit expected quantity of a bidder of group 2 is given by $\lim _{n_{1} \rightarrow \infty} \mathbb{E}\left[x_{2}(t)\right]=\frac{\bar{\theta}_{2}-\bar{\theta}_{1}+\left(\lim _{n \rightarrow \infty} d_{1}+\lambda_{1}\right) q}{\lim _{1} d_{1}+\lambda_{2}}$.

If the small group is fully informed $\left(\sigma_{\varepsilon_{2}}^{2}=0\right)$ and the large group is entirely uninformed $\left(\sigma_{\varepsilon_{1}}^{2} \rightarrow \infty\right)$, then: $\bar{n}_{2}=2 \rho$; an equilibrium always exists for $n_{2}>2$; and the equilibrium coefficients for group 2 are $\lim _{n_{1} \rightarrow \infty} b_{2}=0$, and $\lim _{n_{1} \rightarrow \infty} a_{2}=\lim _{n_{1} \rightarrow \infty} c_{2}=\frac{n_{2}-2 \rho}{\left(n_{2}-\rho\right) \lambda_{2}}$. In this case, the groups' relative market power is given by $\lim _{n_{1} \rightarrow \infty}\left(d_{2} / d_{1}\right)=1+\frac{\rho}{n_{2}-\rho}$. As $\rho$ increases the relative market power of group 2 increases also.

### 4.2 A large price-taking market

Consider now the following setup. There is a continuum of bidders along the interval $[0,1]$, and we let $q$ denote the aggregate (average) quantity supplied in the market. Suppose that a fraction $\mu_{i}\left(0<\mu_{i}<1\right)$ of these bidders are traders of type $i, i=1,2$. Then the following proposition characterizes the equilibrium of this continuum economy and shows that it is the limit of a finite economy's equilibrium.

Proposition 6. Let $Q=\left(n_{1}+n_{2}\right) q$. Suppose that $n_{1}$ and $n_{2}$ both approach to infinity and that $n_{i} /\left(n_{1}+n_{2}\right)$ converges to $\mu_{i}\left(0<\mu_{i}<1\right)$ for $i=1,2$. Then, the equilibrium coefficients converge to the equilibrium coefficients of the equilibrium of the continuum economy setup, which
are given by

$$
\begin{aligned}
b_{i} & =\frac{\widehat{\sigma}_{\varepsilon_{i}}^{2}\left(\rho \lambda_{j} q+\mu_{j}\left(\bar{\theta}_{i}-\rho \bar{\theta}_{j}\right)\right)}{\mu_{i} \rho \lambda_{j} \widehat{\sigma}_{\varepsilon_{i}}^{2}+\mu_{j} \lambda_{i}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)}, a_{i}=\frac{\mu_{j}\left(1-\rho^{2}\right)}{\mu_{i} \rho \lambda_{j} \widehat{\sigma}_{\varepsilon_{i}}^{2}+\mu_{j} \lambda_{i}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)}, \text { and } \\
c_{i} & =\frac{\mu_{j}(1-\rho)\left(1+\rho+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)}{\mu_{i} \rho \lambda_{j} \widehat{\sigma}_{\varepsilon_{i}}^{2}+\mu_{j} \lambda_{i}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)}, \text { where } i, j=1,2, j \neq i
\end{aligned}
$$

## 5 Welfare analysis

This section focuses on the welfare loss at the equilibrium. We characterize the equilibrium and efficient allocations in Subsection 5.1 and analyze deadweight losses in Subsection 5.2.

### 5.1 Characterizing the equilibrium and efficient allocations

Recall that $t_{i}=\mathbb{E}\left[\theta_{i} \mid s_{1}, s_{2}\right], i=1,2$, that is, the predicted values with full information $\left(s_{1}, s_{2}\right)$ and $t=\left(t_{1}, t_{2}\right)$. The strategies in the equilibrium induce outcomes as functions of the realized vector of predicted values $t$ and are given in Equation (5). One can easily show that the equilibrium outcome solves the following distorted benefit maximization program: ${ }^{19}$

$$
\begin{gathered}
\max _{x_{1}, x_{2}} \mathbb{E}\left[n_{1}\left(\theta_{1} x_{1}-\left(d_{1}+\lambda_{1}\right) x_{1}^{2} / 2\right)+n_{2}\left(\theta_{2} x_{2}-\left(d_{2}+\lambda_{2}\right) x_{2}^{2} / 2\right) \mid t\right] \\
\text { s.t. } n_{1} x_{1}+n_{2} x_{2}=Q
\end{gathered}
$$

where $d_{1}$ and $d_{2}$ are the equilibrium parameters. The efficient allocation would obtain if we set $d_{1}=d_{2}=0$, which corresponds to a price-taking equilibrium (denoted by superscript $o$ ). The equilibrium strategy of a type- $i$ bidder $(i=1,2)$ will be of the form $X_{i}^{o}\left(s_{i}, p\right)=b_{i}^{o}+a_{i}^{o} s_{1}-c_{i}^{o} p$, $i=1,2$, and is derived by maximizing the following program:

$$
\max _{x_{i}}\left(\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p\right) x_{i}-\lambda_{i} x_{i}^{2} / 2
$$

while taking prices as given. The FOC of this optimization problem yields

$$
\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p-\lambda_{i} x_{i}=0
$$

After identifying coefficients and solving the corresponding system of equations, we find that there exists a unique equilibrium in this setup. The equilibrium coefficients coincide with those in Proposition 6 for the continuum market.

Proposition 7. Let $Q=\left(n_{1}+n_{2}\right) q$ and let $\mu_{i}=n_{i} /\left(n_{1}+n_{2}\right)$ for $i=1,2$. Then there exists a unique price-taking equilibrium, and the equilibrium coefficients coincide with the equilibrium

[^9]coefficients of the continuum setup (whose expressions are given in the statement of Proposition $6)$.

Our next corollary provides some comparative statics results.
Corollary 3. Let $\rho \sigma_{\varepsilon_{1}}^{2} \sigma_{\varepsilon_{2}}^{2}>0$. Then the only equilibrium coefficients affected by $Q, \bar{\theta}_{i}$, and $\bar{\theta}_{j}$ are the intercepts of the demand functions (with $b_{i}^{o}$ increasing in $\bar{\theta}_{i}$ and $Q$ and decreasing in $\bar{\theta}_{j}$ ) for $i, j=1,2$ and $i \neq j$. Furthermore, the demands of group $i$ are less sensitive to private signals and prices (lower $a_{i}$ and $c_{i}$ ) in response to an increase in $\lambda_{i}, \lambda_{j}, \rho, \sigma_{\varepsilon_{i}}^{2}$, and $\mu_{i}$, and to a decrease in $\mu_{j}$; however, group $i$ 's demands are not affected by $\sigma_{\varepsilon_{j}}^{2}$.

Observe that, under competitive behavior, we can derive an additional comparative statics result: the relationship between the equilibrium coefficients and the proportion of individuals in group 1. In particular, increasing the proportion $\mu_{1}$ of type- 1 traders leads, for those traders, to an increased information-sensitivity weight of the price (higher $\left.\Psi_{1}\left(n_{1} c_{1}^{o}+n_{2} c_{2}^{o}\right)\left(n_{2} a_{2}^{o}\right)^{-1}\right)$ and so a lower overall response to the price $\left(c_{1}^{o}=\lambda_{1}^{-1}\left(1-\Psi_{1}\left(n_{1} c_{1}^{o}+n_{2} c_{2}^{o}\right)\left(n_{2} a_{2}^{o}\right)^{-1}\right)\right)$; the opposite holds for type- 2 traders.

Thus the auction outcome can be obtained as the solution to a maximization problem with a more concave objective function than the expected total surplus, which suggests that inefficiency may be eliminated by quadratic subsidies $\left(\kappa_{i} x_{i}^{2} / 2, i=1,2\right)$ that compensate for the distortions. The per capita subsidy rate $\left(\kappa_{i}\right)$ to a trader of type $i$ must be such that it compensates for the distortion $d_{i}\left(\kappa_{i}\right)$ while accounting for the subsidy. Since the aim is to induce competitive behavior, the trader should be led to respond with $c_{i}^{o}$ to the price. This means that the exact amount of $\kappa_{i}$ must be $d_{i}\left(c_{1}^{o}, c_{2}^{o}\right)$, since that would be the distortion arising when traders use the competitive linear strategies. The following proposition shows that, if subsidies are selected properly, then bidders behave competitively and so the equilibrium allocation is efficient.

Proposition 8. Let $i=1,2$ and $i \neq j$. Then the efficient allocation is induced by the quadratic subsidies $\kappa_{i} x_{i}^{2} / 2$, where $\kappa_{i}=d_{i}\left(c_{i}^{o}, c_{j}^{o}\right)=1 /\left(\left(n_{i}-1\right) c_{i}^{o}+n_{j} c_{j}^{o}\right)$. If $\rho \sigma_{\varepsilon_{i}}^{2} \sigma_{\varepsilon_{j}}^{2}>0$, then the per capita subsidy rates $\left(\kappa_{i}, i=1,2\right)$ increase with $\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}, \lambda_{1}$, and $\lambda_{2}$ but decrease with $n_{1}$ and $n_{2}$. We have that $\kappa_{1}<\kappa_{2}$ iff $c_{1}^{o}<c_{2}^{o}$.

Combining Propositions 7 and 8 now yields closed-form expressions for the optimal subsidy rates:

$$
\kappa_{i}=\frac{1}{n_{j}(1-\rho)}\left(\frac{\left(n_{i}-1\right)\left(1+\widehat{\sigma}_{\varepsilon_{i}}^{2}+\rho\right)}{n_{i} \lambda_{j} \rho \widehat{\sigma}_{\varepsilon_{i}}^{2}+n_{j} \lambda_{i}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)}+\frac{n_{i}\left(1+\widehat{\sigma}_{\varepsilon_{j}}^{2}+\rho\right)}{n_{i} \lambda_{j}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{j}}^{2}\right)+n_{j} \lambda_{i} \rho \widehat{\sigma}_{\varepsilon_{j}}^{2}}\right)^{-1}
$$

$i=1,2, i \neq j$. If $\rho=0$ (or, with full information, if $\sigma_{\varepsilon_{i}}^{2}=0, i=1,2$ ), then $\kappa_{i}^{f}=$ $1 /\left(\left(n_{i}-1\right) \lambda_{i}^{-1}+n_{j} \lambda_{j}^{-1}\right), i=1,2$. Proposition 8 implies that the optimal subsidy rates with
incomplete information and learning from prices are higher than with full information: $\kappa_{i}>\kappa_{i}^{f}$ if (a) $\rho>0$ and (b) at least one of $\widehat{\sigma}_{\varepsilon_{1}}^{2}$ or $\widehat{\sigma}_{\varepsilon_{2}}^{2}$ is strictly positive. If $\sigma_{\varepsilon_{1}}^{2}=\infty$ then $\kappa_{i}, i=1,2$, is independent of $\sigma_{\varepsilon_{2}}^{2}$.

Note that the optimal subsidy rates are decreasing in the number of traders because, when there are many agents, competitive behavior is already being approached in the market without subsidies. Moreover, $\operatorname{sgn}\left\{\kappa_{1}-\kappa_{2}\right\}=\operatorname{sgn}\left\{c_{1}^{o}-c_{2}^{o}\right\}$. Hence $\kappa_{1}<\kappa_{2}$ iff $c_{1}^{o}<c_{2}^{o}$. The implication is that the bidders who require a higher per capita subsidy rate are the ones whose demands are more sensitive to price. Corollary 3 allows us to conclude that, if there is a group with more precise private information, with lower transaction costs, and that is less numerous, then it is the group meriting a higher per capita subsidy rate. The reason is that the stronger group's strategic behavior is more pronounced and so it must receive more compensation in order to become competitive. If $\bar{\theta}_{2} \geq \bar{\theta}_{1}$, then $\mathbb{E}\left[\left(x_{2}^{o}(t)\right)^{2}\right]>\mathbb{E}\left[\left(x_{1}^{o}(t)\right)^{2}\right]$, from which it follows that the bidders from the stronger group (group 2) should receive the higher expected subsidy. ${ }^{20}$ However, if $\bar{\theta}_{1}>\bar{\theta}_{2}$, then there are parameter configurations under which bidders from the weaker group (group 1) should receive the higher expected subsidy, even though $\kappa_{1}<\kappa_{2}$. These conclusions would have to be revised if redistributive considerations come into play. ${ }^{21}$

Our result has policy implications. It implies, for example, that a central bank seeking an efficient distribution of liquidity among banks should relax collateral requirements (i.e., provide a larger subsidy) to the strong group. This prescription sounds counterintuitive because the efficiency motive may conflict with the central bank's function as lender of last resort, which often involves shoring up weak banks (e.g., the European Central Bank relaxing the collateral requirements for Greek banks to avoid a meltdown of that country's banking system). Another example is that of a wholesale electricity market characterized by a small (oligopolistic) group and a fringe; in this case, a regulator looking to improve productive efficiency should set a higher subsidy rate for the oligopolistic group. This could be accomplished by offering differential subsidies to renewable energy technologies, for instance, that lower the marginal cost of production.

[^10]
### 5.2 Deadweight loss

The expected deadweight loss, $\mathbb{E}[D W L]$, at an anonymous allocation $\left(x_{1}(t), x_{2}(t)\right)$ is the difference between expected total surplus at the efficient allocation, ETS ${ }^{o}$, and at the baseline allocation, denoted simply by ETS. Lemma A4 in the Appendix shows that

$$
\mathbb{E}[D W L]=\frac{1}{2} \lambda_{1} n_{1} \mathbb{E}\left[\left(x_{1}(t)-x_{1}^{o}(t)\right)^{2}\right]+\frac{1}{2} \lambda_{2} n_{2} \mathbb{E}\left[\left(x_{2}(t)-x_{2}^{o}(t)\right)^{2}\right],
$$

where $\left(x_{1}^{o}(t), x_{2}^{o}(t)\right)$ corresponds to the price-taking equilibrium. We can use the equilibrium expressions for $\left(x_{1}(t), x_{2}(t)\right)$ and $\left(x_{1}^{o}(t), x_{2}^{o}(t)\right)$ to show that

$$
\begin{aligned}
\mathbb{E}[D W L]= & \frac{n_{2} n_{1}\left(n_{2} d_{1}+n_{1} d_{2}\right)^{2}}{2\left(n_{2} \lambda_{1}+n_{1} \lambda_{2}\right)\left(n_{2}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(d_{2}+\lambda_{2}\right)\right)^{2}} \mathbb{E}\left(t_{1}-t_{2}\right)^{2}+ \\
& +\frac{n_{2} n_{1}\left(n_{2} d_{1}+n_{1} d_{2}\right)\left(\lambda_{2} d_{1}-\lambda_{1} d_{2}\right)}{\left(n_{2} \lambda_{1}+n_{1} \lambda_{2}\right)\left(n_{2}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(d_{2}+\lambda_{2}\right)\right)^{2}} Q\left(\bar{\theta}_{1}-\bar{\theta}_{2}\right)+ \\
& \frac{n_{2} n_{1}\left(\lambda_{1} d_{2}-\lambda_{2} d_{1}\right)^{2}}{2\left(n_{2} \lambda_{1}+n_{1} \lambda_{2}\right)\left(n_{2}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(d_{2}+\lambda_{2}\right)\right)^{2}} Q^{2},
\end{aligned}
$$

where $\mathbb{E}\left[\left(t_{1}-t_{2}\right)^{2}\right]=\left(\bar{\theta}_{1}-\bar{\theta}_{2}\right)^{2}+(1-\rho)^{2} \sigma_{\theta}^{2}\left(2(1+\rho)+\widehat{\sigma}_{\varepsilon_{1}}^{2}+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right) /\left(\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}\right)$.
The expected deadweight loss consists of three terms. The first term is the only one present in a double auction $(Q=0)$. This term, which is due to uncertainty and information, is the product of two factors. One factor

$$
\frac{n_{2} n_{1}\left(n_{2} d_{1}+n_{1} d_{2}\right)^{2}}{2\left(n_{2} \lambda_{1}+n_{1} \lambda_{2}\right)\left(n_{2}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(d_{2}+\lambda_{2}\right)\right)^{2}}
$$

increases with $d_{1}$ and $d_{2}$. Since $d_{1}$ and $d_{2}$ are each increasing in $\rho$, it follows that this multiplier also increases with $\rho$. The other factor, $\mathbb{E}\left[\left(t_{1}-t_{2}\right)^{2}\right]$, decreases with $\rho$ and with $\widehat{\sigma}_{\varepsilon_{i}}^{2}$ and increases with $\left(\bar{\theta}_{1}-\bar{\theta}_{2}\right)^{2}$; it vanishes when $\rho=1$ or when there is no uncertainty $\left(\sigma_{\theta}^{2}=0\right)$ provided that $\bar{\theta}_{1}=\bar{\theta}_{2}$. As a result, the first term of $\mathbb{E}[D W L]$ may be either increasing or decreasing in $\rho$. Note that if $\sigma_{\varepsilon_{1}}^{2}=\infty$, then $d_{1}$ and $d_{2}$ are independent of $\sigma_{\varepsilon_{2}}^{2}$, and $\mathbb{E}[D W L]$ decreases with $\sigma_{\varepsilon_{2}}^{2}$.

The third term derives from the absorption of $Q$ by the traders, and it increases with the quantity offered $(Q)$ as well as with the difference (in absolute terms) between $d_{1} / d_{2}$ and $\lambda_{1} / \lambda_{2}$. The second term is an interaction term that is positive for $Q>0$ iff $\left(\lambda_{2} d_{1}-\lambda_{1} d_{2}\right)\left(\bar{\theta}_{1}-\bar{\theta}_{2}\right)>0$, that is, when the relative distortion between groups $\left(d_{i} / d_{j}\right)$ is large whenever $\bar{\theta}_{i}>\bar{\theta}_{j}$. When $d_{1} / d_{2}=\lambda_{1} / \lambda_{2}$, the expected deadweight loss consists of the first term only. That is because, in this case, the non-informational trading term corresponding to the equilibrium with imperfect competition coincides with the one corresponding to the competitive equilibrium. Note that if we interpret the traders as producers competing to supply a fixed demand $Q$, then the condition
$d_{1} / d_{2}=\lambda_{1} / \lambda_{2}$ means that the ratio of the production of the two types of firms is aligned with the ratio of the slopes of their respective marginal costs. This condition guarantees productive efficiency provided that $\bar{\theta}_{1}=\bar{\theta}_{2}$ and $\rho=1$ and, since demand is fixed, this coincides with overall efficiency.

Furthermore, if group 1 has higher transaction costs $\left(\lambda_{1}>\lambda_{2}\right)$, is more numerous $\left(n_{1}>n_{2}\right)$, and is less informed $\left(\sigma_{\varepsilon_{1}}^{2}>\sigma_{\varepsilon_{2}}^{2}\right)$ than group 2, then $d_{1} / d_{2}<\lambda_{1} / \lambda_{2}$ (and therefore $\lambda_{2} d_{1}-\lambda_{1} d_{2}<$ $0)$. In this case, the third term in our expression for $\mathbb{E}[D W L]$ is not null. In addition, if group 1 ex ante values the asset less $\left(\bar{\theta}_{1} \leq \bar{\theta}_{2}\right)$, then the second term in our expression for $\mathbb{E}[D W L]$ is positive. The expected deadweight loss increases with $Q$ and $\left|\bar{\theta}_{1}-\bar{\theta}_{2}\right|$ when the stronger group values the asset no less than does the weaker group.

Under full information (i.e., $\sigma_{\varepsilon_{1}}^{2}=\sigma_{\varepsilon_{2}}^{2}=0$ ), both $d_{1}$ and $d_{2}$ are independent of $\rho$; in this case, then $\mathbb{E}[D W L]$ decreases with $\rho$. Similarly, if $\rho=0$, then $d_{1}$ and $d_{2}$ are independent of $\sigma_{\varepsilon_{1}}^{2}$ and $\sigma_{\varepsilon_{2}}^{2}$, from which it follows that $\mathbb{E}[D W L]$ decreases with $\sigma_{\varepsilon_{1}}^{2}$ and $\sigma_{\varepsilon_{2}}^{2}$. Some of these results are summarized in our last proposition.

## Proposition 9.

(i) The expected deadweight loss may be increasing or decreasing in the information parameters $\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}\right.$, and $\left.\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)$ and, therefore, market power $\left(d_{1}, d_{2}\right)$ and the $\mathbb{E}[D W L]$ may be negatively associated.
(ii) The $\mathbb{E}[D W L]$ increases with payoff asymmetry and with $Q$ whenever the stronger group (say group 2, with $\lambda_{1}>\lambda_{2}, n_{1}>n_{2}$, and $\sigma_{\varepsilon_{1}}^{2}>\sigma_{\varepsilon_{2}}^{2}$ ) values the asset no less than does the weaker group (i.e., $\bar{\theta}_{1} \leq \bar{\theta}_{2}$ ).
(iii) When groups are symmetric, the expected deadweight loss is independent of $Q$, and market power $d$ and the $\mathbb{E}[D W L \mid t]$ are positively associated, given predicted values $t$, for changes in information parameters. This need not be the case with asymmetric groups (e.g., for $Q$ large, $d_{i} / d_{j}>\lambda_{i} / \lambda_{j}$ implies that $\mathbb{E}[D W L \mid t]$ increases in $d_{i}$ and decreases in $\left.d_{j}\right)$.

## 6 Concluding remarks

We analyze a divisible good uniform-price auction, where two types of bidders compete. Each of these two groups contains a finite number of identical bidders. At the unique equilibrium, a group's relative market power increases with the precision of private information and decreases with the group's transaction costs. Consistently with the empirical evidence, we find that an increase in the transaction costs of a group of bidders induces a strategic response from the other group, whose members then submit steeper schedules. The group that is stronger (because it has more precise private information, faces lower transaction costs, and is more oligopsonistic)
has more market power and must therefore receive a higher subsidy to behave competitively. The expected deadweight loss increases with the quantity auctioned and with the degree of payoff asymmetry provided the stronger group values the asset no less than does the weaker group.

Our findings have policy implications. Consider a regulator who wants to reduce inefficiency in an industry with two groups of firms (e.g., a small oligopolistic group and a competitive fringe). This regulator must bear in mind that any intervention directed toward one group will also affect the other's behavior. In addition, the regulator should set a higher subsidy rate for the group that has better information, is more oligopsonistic, and has lower transaction costs. The framework developed here can be adapted to study competition policy analyzing the effects of merger and industry capacity redistribution.

## Appendix

Proposition 1 will follow from Lemmata A1 and A2.
Lemma A1. Let $\rho<1$. In equilibrium, the demand function for a trader of type $i, i=1,2$, is given by $X_{i}\left(s_{i}, p\right)=\left(\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p\right) /\left(d_{i}+\lambda_{i}\right)$, with $d_{i}+\lambda_{i}>0$. The equilibrium coefficients satisfy the following system of equations:

$$
\begin{align*}
b_{i} & =\left(\left(1-\Xi_{i}\right) \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}-\frac{\Psi_{i}\left(n_{i} b_{i}+n_{j} b_{j}-Q\right)}{n_{j} a_{j}}\right) /\left(d_{i}+\lambda_{i}\right)  \tag{9}\\
a_{i} & =\left(\Xi_{i}-\frac{n_{i} a_{i}}{n_{j} a_{j}} \Psi_{i}\right) /\left(d_{i}+\lambda_{i}\right), \text { and }  \tag{10}\\
c_{i} & =\left(1-\frac{\Psi_{i}\left(n_{i} c_{i}+n_{j} c_{j}\right)}{n_{j} a_{j}}\right) /\left(d_{i}+\lambda_{i}\right) \tag{11}
\end{align*}
$$

where $i, j=1,2, j \neq i$. Moreover, in equilibrium, $a_{i}>0, i=1,2$.
Proof: Consider a trader of type $i$. Recall that at the beginning of Subsection 3.1 we obtain $X_{i}\left(s_{i}, p\right)=\left(\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p\right) /\left(d_{i}+\lambda_{i}\right)$ and $\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]=\mathbb{E}\left[\theta_{i} \mid s_{i}, s_{j}\right]$. Since we are looking for strategies of the form $X_{i}\left(s_{i}, p\right)=b_{i}+a_{i} s_{i}-c_{i} p$, from the market clearing condition we have $p=\left(n_{i}\left(b_{i}+a_{i} s_{i}\right)+n_{j}\left(b_{j}+a_{j} s_{j}\right)-Q\right) /\left(n_{i} c_{i}+n_{j} c_{j}\right)$ and, hence,

$$
s_{j}=\left(\left(n_{i} c_{i}+n_{j} c_{j}\right) p+Q-n_{i}\left(b_{i}+a_{i} s_{i}\right)-n_{j} b_{j}\right) /\left(n_{j} a_{j}\right) .
$$

Thus, from Expression (3), it follows that
$\mathbb{E}\left[\theta_{i} \mid s_{i}, s_{j}\right]=\left(1-\Xi_{i}\right) \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}+\Psi_{i}\left(\frac{Q-n_{i} b_{i}-n_{j} b_{j}}{n_{j} a_{j}}\right)+\left(\Xi_{i}-\frac{n_{i} a_{i}}{n_{j} a_{j}} \Psi_{i}\right) s_{i}+\Psi_{i}\left(\frac{n_{i} c_{i}+n_{j} c_{j}}{n_{j} a_{j}}\right) p$.

Substituting the foregoing expression in (1), and then identifying coefficients, we obtain the expressions for the demand coefficients given in (9)-(11).

Finally, we show the positiveness of the coefficients $a_{i}, i=1,2$. From Expression (10), we have $a_{i}=\Xi_{i} /\left(d_{i}+\lambda_{i}+n_{i} \Psi_{i} /\left(n_{j} a_{j}\right)\right), i, j=1,2, j \neq i$. Combining the previous expressions, we have that

$$
\begin{equation*}
a_{i}=\frac{n_{j}\left(\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}\right)}{n_{i} \Psi_{i}\left(d_{j}+\lambda_{j}\right)+\Xi_{j} n_{j}\left(d_{i}+\lambda_{i}\right)}, i, j=1,2, j \neq i \tag{12}
\end{equation*}
$$

Direct computation yields $\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}=\left(1-\rho^{2}\right) /\left(\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}\right)>0$, whenever $\rho<1$. Moreover, using the positiveness of $d_{i}+\lambda_{i}, \Xi_{i}$, and $\Psi_{i}, i=1,2$, we conclude that, in equilibrium, the coefficients $a_{i}, i=1,2$, are strictly positive.

Lemma A2. Let $z=c_{1} / c_{2}$. In equilibrium,

$$
\begin{align*}
b_{i} & =\frac{\Psi_{i}}{n_{i} n_{j}} \frac{n_{i} \Xi_{j} \frac{a_{i}}{a_{j}}-n_{j} \Psi_{j}}{\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}} Q+a_{i}\left(\frac{\Xi_{j} \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}}{\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}}-\bar{\theta}_{i}\right),  \tag{13}\\
a_{i} & =\Delta_{i} c_{i},  \tag{14}\\
c_{1} & =\left(\Xi_{1} \Delta_{1}^{-1}-\frac{n_{1}}{n_{2}}\left(1-\Xi_{1} \Delta_{1}^{-1}\right) z-\frac{z}{\left(n_{1}-1\right) z+n_{2}}\right) / \lambda_{1}, \text { and }  \tag{15}\\
c_{2} & =\left(\Xi_{2} \Delta_{2}^{-1}-\frac{n_{2}}{n_{1}}\left(1-\Xi_{2} \Delta_{2}^{-1}\right) \frac{1}{z}-\frac{1}{n_{1} z+n_{2}-1}\right) / \lambda_{2}, \tag{16}
\end{align*}
$$

where $\Delta_{i}=\left(\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}\right) /\left(\Xi_{j}-\Psi_{i}\right)=1 /\left(1+(1+\rho)^{-1} \widehat{\sigma}_{\varepsilon_{i}}^{2}\right), i, j=1,2, j \neq i$. Moreover, $z$ is the unique positive solution to the cubic polynomial $p(z)=p_{3} z^{3}+p_{2} z^{2}+p_{1} z+p_{0}$, with

$$
\begin{aligned}
p_{3}= & n_{1}^{2}\left(n_{1}-1\right)\left(n_{2} \Xi_{2} \Delta_{2}^{-1} \lambda_{1}+n_{1}\left(1-\Xi_{1} \Delta_{1}^{-1}\right) \lambda_{2}\right), \\
p_{2}= & n_{1}\left(\left(3 n_{2} n_{1}-n_{1}-2 n_{2}+1\right)\left(n_{2} \Xi_{2} \Delta_{2}^{-1} \lambda_{1}-n_{1} \Xi_{1} \Delta_{1}^{-1} \lambda_{2}\right)+\right. \\
& \left.+\lambda_{2} n_{1}\left(2 n_{2} n_{1}-n_{1}+1\right)-\left(n_{1}-1\right)\left(n_{2}+1\right) n_{2} \lambda_{1}\right), \\
p_{1}= & n_{2}\left(\left(3 n_{2} n_{1}-2 n_{1}-n_{2}+1\right)\left(n_{2} \Xi_{2} \Delta_{2}^{-1} \lambda_{1}-n_{1} \Xi_{1} \Delta_{1}^{-1} \lambda_{2}\right)+\right. \\
& \left.+\lambda_{2} n_{1}\left(n_{2}-1\right)\left(n_{1}+1\right)-\left(2 n_{2} n_{1}-n_{2}+1\right) n_{2} \lambda_{1}\right), \text { and } \\
p_{0}= & -n_{2}^{2}\left(n_{2}-1\right)\left(n_{2}\left(1-\Xi_{2} \Delta_{2}^{-1}\right) \lambda_{1}+n_{1} \Xi_{1} \Delta_{1}^{-1} \lambda_{2}\right) .
\end{aligned}
$$

Proof: In relation to the expression for $b_{i}$, notice that (10) implies

$$
\begin{equation*}
d_{i}+\lambda_{i}=\left(\Xi_{i}-\frac{n_{i} a_{i}}{n_{j} a_{j}} \Psi_{i}\right) / a_{i}, i, j=1,2, j \neq i \tag{17}
\end{equation*}
$$

Substituting these expressions in (9), it follows that

$$
\begin{equation*}
b_{i}=a_{i} \frac{\left(1-\Xi_{i}\right) \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}-\frac{\Psi_{i}\left(n_{i} b_{i}+n_{j} b_{j}-Q\right)}{n_{j} a_{j}}}{\Xi_{i}-\frac{n_{i} a_{i}}{n_{j} a_{j}} \Psi_{i}}, i, j=1,2, j \neq i . \tag{18}
\end{equation*}
$$

Thus,
$n_{i} b_{i}+n_{j} b_{j}=n_{i} a_{i} \frac{\left(1-\Xi_{i}\right) \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}-\frac{\Psi_{i}\left(n_{i} b_{i}+n_{j} b_{j}-Q\right)}{n_{j} a_{j}}}{\Xi_{i}-\frac{n_{i} a_{i}}{n_{j} a_{j}} \Psi_{i}}+n_{j} a_{j} \frac{\left(1-\Xi_{j}\right) \bar{\theta}_{j}-\Psi_{j} \bar{\theta}_{i}-\frac{\Psi_{j}\left(n_{i} b_{i}+n_{j} b_{j}-Q\right)}{n_{i} a_{i}}}{\Xi_{j}-\frac{n_{j} a_{j}}{n_{i} a_{i}} \Psi_{j}}$,
Isolating $n_{i} b_{i}+n_{j} b_{j}$ in the previous formula and substituting the resulting expression in (18), Expression (13) is obtained.

Concerning the expression for $a_{i}$, substituting (17) in (11), it follows that

$$
\begin{equation*}
c_{i}=a_{i}\left(1-\frac{\Psi_{i}\left(n_{i} c_{i}+n_{j} c_{j}\right)}{n_{j} a_{j}}\right) /\left(\Xi_{i}-\frac{n_{i} a_{i}}{n_{j} a_{j}} \Psi_{i}\right), i, j=1,2, j \neq i \tag{19}
\end{equation*}
$$

Hence, $n_{i} c_{i}+n_{j} c_{j}=n_{i} a_{i} \frac{n_{j} a_{j}-\Psi_{i}\left(n_{i} c_{i}+n_{j} c_{j}\right)}{n_{j} a_{j} \Xi_{i}-n_{i} a_{i} \Psi_{i}}+n_{j} a_{j} \frac{n_{i} a_{i}-\Psi_{j}\left(n_{i} c_{i}+n_{j} c_{j}\right)}{n_{i} a_{i} \Xi_{j}-n_{j} a_{j} \Psi_{j}}$, which implies that $n_{i} c_{i}+n_{j} c_{j}=$ $\left(n_{i} a_{i}\left(\Xi_{j}-\Psi_{i}\right)+n_{j} a_{j}\left(\Xi_{i}-\Psi_{j}\right)\right) /\left(\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}\right)$. Then, substituting the previous expression in Equation (19), we obtain $c_{i}=\left(\Xi_{j}-\Psi_{i}\right) a_{i} /\left(\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}\right)$, which is equivalent to (14).

In relation to $c_{1}$ and $c_{2}$, using the expression for $d_{i}$ and (14), (17) implies that

$$
\lambda_{i}=\left(\frac{\Xi_{i}}{\Delta_{i}}-\frac{n_{i} \Psi_{i} c_{i}}{n_{j} \Delta_{j} c_{j}}\right) c_{i}^{-1}-\left(\left(n_{i}-1\right) c_{i}+n_{j} c_{j}\right)^{-1}, i, j=1,2, j \neq i
$$

or, since

$$
\begin{equation*}
\Psi_{i} \Delta_{j}^{-1}=1-\Xi_{i} \Delta_{i}^{-1} \tag{20}
\end{equation*}
$$

$\lambda_{i}=\left(\Xi_{i} \Delta_{i}^{-1}-\frac{n_{i}}{n_{j}}\left(1-\Xi_{i} \Delta_{i}^{-1}\right) \frac{c_{i}}{c_{j}}\right) c_{i}^{-1}-\left(\left(n_{i}-1\right) c_{i}+n_{j} c_{j}\right)^{-1}, i, j=1,2, j \neq i$, which imply (15) and (16) since $z=c_{1} / c_{2}$. Moreover, dividing the previous two equalities, it follows that

$$
\begin{equation*}
\frac{\lambda_{1}}{\lambda_{2}}=\frac{\Xi_{1} \Delta_{1}^{-1}-\frac{n_{1}}{n_{2}}\left(1-\Xi_{1} \Delta_{1}^{-1}\right) z-z\left(\left(n_{1}-1\right) z+n_{2}\right)^{-1}}{\Xi_{2} \Delta_{2}^{-1} z-\frac{n_{2}}{n_{1}}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)-z\left(n_{1} z+n_{2}-1\right)^{-1}} \tag{21}
\end{equation*}
$$

After some algebra, (21) is equivalent to $p(z)=0$, where $p(z)=p_{3} z^{3}+p_{2} z^{2}+p_{1} z+p_{0}$, as stated in this lemma. Notice that $p(0)<0$ and $\lim _{z \rightarrow \infty} p(z)=\infty$. Consequently, there exists $z \in(0, \infty)$ such that $p(z)=0$. Furthermore, we have that $p_{2} / n_{1}>p_{1} / n_{2}$. This property implies that there exists only one sign change in the coefficients of $p(z)$. Applying the Descartes' rule, we conclude that there exists a unique positive root of $p(z)$.

Proof of Proposition 2: (Necessity). From Proposition 1 we know that $a_{i}>0, i=1,2$. Combining this property with expressions given in (14), we have that, in equilibrium, the coefficients $c_{1}$ and $c_{2}$ are strictly positive. Moreover, (15) and (16) can be rewritten as

$$
c_{1}=\frac{q_{N}(z)}{\left(\left(n_{1}-1\right) z+n_{2}\right) n_{2} \lambda_{1}} \text { and } c_{2}=\frac{q_{D}(z)}{\left(n_{1} z+n_{2}-1\right) n_{1} z \lambda_{2}},
$$

where $q_{N}(z)=n_{2}^{2} \Xi_{1} \Delta_{1}^{-1}+n_{2}\left(\Xi_{1} \Delta_{1}^{-1}\left(2 n_{1}-1\right)-\left(n_{1}+1\right)\right) z-\left(n_{1}-1\right)\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} z^{2}$ and $q_{D}(z)=-n_{2}\left(n_{2}-1\right)\left(1-\Xi_{2} \Delta_{2}^{-1}\right)+n_{1}\left(\Xi_{2} \Delta_{2}^{-1}\left(2 n_{2}-1\right)-\left(n_{2}+1\right)\right) z+n_{1}^{2} \Xi_{2} \Delta_{2}^{-1} z^{2}$.

Let $\bar{z}_{N}$ and $\bar{z}_{D}$ denote the highest root of $q_{N}(z)$ and $q_{D}(z)$, respectively. Notice that the positiveness of $c_{1}$ and $c_{2}$ is equivalent to $\bar{z}_{N}>z>\bar{z}_{D}$. Therefore, $\bar{z}_{N}>\bar{z}_{D}$.
(Sufficiency). Suppose that $\bar{z}_{N}>\bar{z}_{D}$. Recall that Lemma A2 shows that there exists a unique positive value of $z$ that solves (21), which can be rewritten as $\frac{\lambda_{1}}{\lambda_{2}}=\frac{n_{1}\left(n_{2}-1+n_{1} z\right) q_{N}(z)}{\left(n_{2}+\left(n_{1}-1\right) z\right) n_{2} q_{D}(z)}$. This implies that $\bar{z}_{N}>z>\bar{z}_{D}$. Notice that these inequalities guarantee the positiveness of $c_{1}$ and $c_{2}$. Therefore, $d_{1}$ and $d_{2}$ are strictly positive, and consequently, the SOC of the optimization problems are satisfied. Thus, we conclude that whenever $\bar{z}_{N}>\bar{z}_{D}$ there exists a unique equilibrium.

Proof of Corollary 1: Notice that if $n_{2}=1$, then $\bar{z}_{D}=\left(2 \Delta_{2} \Xi_{2}^{-1}-1\right) / n_{1}$. Thus, the condition that guarantees the existence of an equilibrium is equivalent to $q_{N}\left(\left(2 \Delta_{2} \Xi_{2}^{-1}-1\right) / n_{1}\right)>0$. Direct computation yields that the last inequality holds iff

$$
\left(2-\Xi_{2} \Delta_{2}^{-1}\right)\left(1-\Xi_{1} \Delta_{1}^{-1}-\Xi_{2} \Delta_{2}^{-1}\right)+\left(\Xi_{2} \Delta_{2}^{-1}-2\left(1-\Xi_{1} \Delta_{1}^{-1}\right)\right) n_{1}>0
$$

As $1<\Xi_{1} \Delta_{1}^{-1}+\Xi_{2} \Delta_{2}^{-1}$, we conclude that an equilibrium exists iff $\Xi_{2} \Delta_{2}^{-1}>2\left(1-\Xi_{1} \Delta_{1}^{-1}\right)$ and $n_{1}>\left(2-\Xi_{2} \Delta_{2}^{-1}\right)\left(\Xi_{1} \Delta_{1}^{-1}+\Xi_{2} \Delta_{2}^{-1}-1\right) /\left(\Xi_{2} \Delta_{2}^{-1}-2\left(1-\Xi_{1} \Delta_{1}^{-1}\right)\right)$. Using the expressions of $\Xi_{i}$ and $\Delta_{i}$, the previous two inequalities are equivalent to $1-\rho^{2}>(2 \rho-1) \widehat{\sigma}_{\varepsilon_{1}}^{2}$ and $n_{1}>$ $\bar{n}_{1}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)$, where

$$
\begin{equation*}
\bar{n}_{1}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)=1+\frac{\rho\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left((1+\rho)\left(\widehat{\sigma}_{\varepsilon_{1}}^{2}+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)+2 \widehat{\sigma}_{\varepsilon_{1}}^{2} \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)}{(1+\rho)\left(\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}\right)\left(1-\rho^{2}-(2 \rho-1) \widehat{\sigma}_{\varepsilon_{1}}^{2}\right)} . \tag{22}
\end{equation*}
$$

It can be shown that $\bar{n}_{1}$ increases with $\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}$, and $\widehat{\sigma}_{\varepsilon_{2}}^{2}$. In particular, if $\sigma_{\varepsilon_{2}}^{2}=0$, then (22) can be rewritten as $\bar{n}_{1}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, 0\right)=1+\rho \widehat{\sigma}_{\varepsilon_{1}}^{2} /\left(1-\rho^{2}-(2 \rho-1) \widehat{\sigma}_{\varepsilon_{1}}^{2}\right)$.

Consider the case that one group is formed by a unique trader perfectly informed and with no transaction $\operatorname{costs}\left(n_{2}=1, \lambda_{2}=0\right.$ and $\left.\sigma_{\varepsilon_{2}}^{2}=0\right)$. Then, $z=\bar{z}_{D}=1 / n_{1}, \Xi_{1}=$ $\left(1-\rho^{2}\right) /\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right), \Psi_{1}=\rho \widehat{\sigma}_{\varepsilon_{1}}^{2} /\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right), \Xi_{2}=1$, and $\Psi_{2}=0$. From Lemma A2, the coefficients of the demand functions are given by: $b_{1}=\frac{\rho \widehat{\varepsilon}_{\varepsilon_{1}}^{2}}{n_{1}(1-\rho)\left(1+\rho+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)} Q+a_{1} \frac{\widehat{\sigma}_{\varepsilon_{1}}^{2}}{1-\rho^{2}}\left(\bar{\theta}_{1}-\rho \bar{\theta}_{2}\right)$, $a_{1}=\Delta_{1} c_{1}, c_{1}=\frac{2\left(n_{1}-\bar{n}_{1}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)\right)\left(1-\rho^{2}-\widehat{\sigma}_{\varepsilon_{1}}^{2}(2 \rho-1)\right)}{\lambda_{1}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(2 n_{1}-1\right)}, b_{2}=0$, and $a_{2}=c_{2}=\frac{c_{1}}{z}=n_{1} c_{1}$.

Proof of Proposition 3: In what follows we prove the following comparative statics results for $i, j=1,2, i \neq j$ :
a) $\frac{\partial b_{i}}{\partial \bar{\theta}_{i}}>0, \frac{\partial a_{i}}{\partial \bar{\theta}_{i}}=0$ and $\frac{\partial c_{i}}{\partial \bar{\theta}_{i}}=0$,
b) $\frac{\partial b_{i}}{\partial \bar{\theta}_{j}}<0, \frac{\partial a_{i}}{\partial \bar{\theta}_{j}}=0$ and $\frac{\partial c_{i}}{\partial \bar{\theta}_{j}}=0$,
c) $\frac{\partial b_{i}}{\partial Q}>0, \frac{\partial a_{i}}{\partial Q}=0$ and $\frac{\partial c_{i}}{\partial Q}=0$,
d) $\frac{\partial a_{i}}{\partial \lambda_{i}}<0$ and $\frac{\partial c_{i}}{\partial \lambda_{i}}<0$,
e) $\frac{\partial a_{i}}{\partial \lambda_{j}}<0$ and $\frac{\partial c_{i}}{\partial \lambda_{j}}<0$,
f) $\frac{\partial a_{i}}{\partial \rho}<0$ and $\frac{\partial c_{i}}{\partial \rho}<0$,
g) $\frac{\partial}{\partial \sigma_{\varepsilon_{i}}^{2}}\left(d_{i} / d_{j}\right)<0, \frac{\partial}{\partial \sigma_{\varepsilon_{j}}^{2}}\left(d_{i} / d_{j}\right)>0, \frac{\partial}{\partial \lambda_{i}}\left(d_{i} / d_{j}\right)<0$, and $\frac{\partial}{\partial \lambda_{j}}\left(d_{i} / d_{j}\right)>0$,
h) $\frac{\partial a_{i}}{\partial \sigma_{\varepsilon_{i}}}<0$ and $\frac{\partial c_{i}}{\partial \sigma_{\varepsilon_{i}}}<0$,
i) $\frac{\partial a_{i}}{\partial \sigma_{\varepsilon_{j}}^{\varepsilon_{j}}}<0$ and $\frac{\partial c_{i}}{\partial \sigma_{\varepsilon_{j}}^{\varepsilon_{j}}}<0$, and
k) $\frac{\partial d_{i}}{\partial n_{i}}<0$ and $\frac{\partial d_{j}}{\partial n_{i}}<0$.

From Lemma A1, we know that the equilibrium coefficients that depend on $\bar{\theta}_{i}, \bar{\theta}_{j}$ and $Q$ are $b_{1}$ and $b_{2}$. Using Lemma A2 and after some algebra, the results given in a), b) and c) are obtained. In what follows, without any loss of generality, let $i=1$. First, we prove that $\frac{\partial z}{\partial \lambda_{1}}<0$. From Lemma A2, we know that $z$ is the unique positive solution of the following equation:

$$
\begin{equation*}
\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}=0 \tag{23}
\end{equation*}
$$

where $N(z)=\Xi_{1} \Delta_{1}^{-1}-\frac{n_{1}}{n_{2}}\left(1-\Xi_{1} \Delta_{1}^{-1}\right) z-z\left(\left(n_{1}-1\right) z+n_{2}\right)^{-1}$ and

$$
D(z)=\Xi_{2} \Delta_{2}^{-1} z-\frac{n_{2}}{n_{1}}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)-z\left(n_{1} z+n_{2}-1\right)^{-1}
$$

with $\Xi_{i} \Delta_{i}^{-1}=\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{j}}^{2}\right)\left(1+\rho+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)\left(\left(\left(1+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{j}}^{2}\right)-\rho^{2}\right)(1+\rho)\right)^{-1}, i, j=1,2, i \neq$ $j$. Applying the Implicit Function Theorem, $\frac{\partial z}{\partial \lambda_{i}}=-\left(\frac{\partial}{\partial \lambda_{i}}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)\right) / \frac{\partial}{\partial z}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)$, $i=1,2$. As $\frac{\partial}{\partial \lambda_{1}}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)=\frac{1}{\lambda_{2}}>0, \frac{\partial}{\partial \lambda_{2}}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)=-\frac{\lambda_{1}}{\lambda_{2}^{2}}<0$, and $\frac{\partial}{\partial z}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)>0$, because of $z \in\left(\bar{z}_{D}, \bar{z}_{N}\right)$, we conclude that $\frac{\partial z}{\partial \lambda_{1}}<0$ and $\frac{\partial z}{\partial \lambda_{2}}>0$.

Next, we study the relationship between $c^{\prime} s$ and $\lambda_{1}$. Differentiating (16), we have

$$
\frac{\partial c_{2}}{\partial \lambda_{1}}=\frac{\partial c_{2}}{\partial z} \frac{\partial z}{\partial \lambda_{1}}=\frac{1}{\lambda_{2}}\left(\frac{n_{2}}{n_{1}}\left(1-\Xi_{2} \Delta_{2}^{-1}\right) \frac{1}{z^{2}}+\frac{n_{1}}{\left(n_{2}+n_{1} z-1\right)^{2}}\right) \frac{\partial z}{\partial \lambda_{1}}<0
$$

since $\frac{\partial z}{\partial \lambda_{1}}<0$. Moreover, as $c_{1}=z c_{2}$, it follows that $\frac{\partial c_{1}}{\partial \lambda_{1}}=\frac{\partial z}{\partial \lambda_{1}} c_{2}+z \frac{\partial c_{2}}{\partial \lambda_{1}}<0$, because of the positiveness of $c_{2}$ and $z$, and the negativeness of $\frac{\partial z}{\partial \lambda_{1}}$ and $\frac{\partial c_{2}}{\partial \lambda_{1}}$. In relation to $a_{1}$ and $a_{2}$, from (14), direct computation yields $\frac{\partial a_{1}}{\partial \lambda_{1}}<0$ and $\frac{\partial a_{2}}{\partial \lambda_{1}}<0$, since $\frac{\partial c_{1}}{\partial \lambda_{1}}<0$ and $\frac{\partial c_{2}}{\partial \lambda_{1}}<0$.

Now, we study how the correlation coefficient $\rho$ affects $a_{1}$. Let $y=a_{1} / a_{2}$. As $a_{1}=\Delta_{1} c_{1}$ and $a_{2}=\Delta_{2} c_{2}$, then $z=\Delta_{2} y / \Delta_{1}$. Substituting this expression in (21), and after some algebra, we have that

$$
\begin{equation*}
\frac{\lambda_{1}}{\lambda_{2}} y=\frac{\widetilde{N}(y, \rho)}{\widetilde{D}(y, \rho)} \tag{24}
\end{equation*}
$$

where $\tilde{N}(y, \rho)=\frac{1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{2}}^{2}-\frac{n_{1}}{n_{1}} \widehat{\sigma}_{\varepsilon_{1}}^{2} \rho y}{\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}}-\left(\left(n_{1}-1\right) \frac{1+\rho+\widehat{\sigma}_{\varepsilon_{1}}^{2}}{1+\rho}+n_{2} \frac{1+\rho+\widehat{\sigma}_{\varepsilon_{2}}^{2}}{1+\rho} \frac{1}{y}\right)^{-1}$ and

$$
\widetilde{D}(y, \rho)=\frac{1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{1}}^{2}-\frac{n_{2}}{n_{1}} \widehat{\sigma}_{\varepsilon_{2}}^{2} \rho \frac{1}{y}}{\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}}-\left(n_{1} \frac{1+\rho+\widehat{\sigma}_{\varepsilon_{1}}^{2}}{1+\rho} y+\left(n_{2}-1\right) \frac{1+\rho+\widehat{\sigma}_{\varepsilon_{2}}^{2}}{1+\rho}\right)^{-1}
$$

Moreover, $a_{1}=\widetilde{N}(y, \rho) / \lambda_{1}$ and $a_{2}=\widetilde{D}(y, \rho) / \lambda_{2}$. Hence, $\frac{\partial a_{1}}{\partial \rho}=\left(\frac{\partial}{\partial y} \widetilde{N}(y, \rho) \frac{\partial y}{\partial \rho}+\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)\right) / \lambda_{1}$.
Thus, in order to show $\frac{\partial a_{1}}{\partial \rho}<0$, it suffices to prove that

$$
\begin{equation*}
\frac{\partial}{\partial y} \widetilde{N}(y, \rho) \frac{\partial y}{\partial \rho}+\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)<0 \tag{25}
\end{equation*}
$$

Direct computation yields $\frac{\partial}{\partial y} \widetilde{N}(y, \rho)<0$. Then, (25) is equivalent to

$$
\begin{equation*}
\frac{\partial y}{\partial \rho}>-\frac{\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)}{\frac{\partial}{\partial y} \widetilde{N}(y, \rho)} \tag{26}
\end{equation*}
$$

Moreover, recall that $y$ in equilibrium is the unique positive value that satisfies (24). Thus, applying the Implicit Function Theorem, it follows that

$$
\frac{\partial y}{\partial \rho}=-\frac{\frac{\partial}{\partial \rho}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\tilde{N}(y, \rho)}{\tilde{D}(y, \rho)}\right)}{\frac{\partial}{\partial y}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\tilde{N}(y, \rho)}{\tilde{D}(y, \rho)}\right)} .
$$

Then, (26) can be rewritten as

$$
-\frac{\frac{\partial}{\partial \rho}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\tilde{N}(y, \rho)}{\widetilde{D}(y, \rho)}\right)}{\frac{\partial}{\partial y}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\widetilde{N}(y, \rho)}{\widetilde{D}(y, \rho)}\right)}>-\frac{\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)}{\frac{\partial}{\partial y} \widetilde{N}(y, \rho)},
$$

or using the fact that in equilibrium $\frac{\partial}{\partial y}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\tilde{N}(y, \rho)}{\tilde{D}(y, \rho)}\right)>0,(26)$ is satisfied iff

$$
\begin{equation*}
-\frac{\partial}{\partial \rho}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\widetilde{N}(y, \rho)}{\widetilde{D}(y, \rho)}\right)>-\frac{\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)}{\frac{\partial}{\partial y} \widetilde{N}(y, \rho)} \frac{\partial}{\partial y}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\widetilde{N}(y, \rho)}{\widetilde{D}(y, \rho)}\right) . \tag{27}
\end{equation*}
$$

Notice that

$$
\frac{\partial}{\partial \rho}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\widetilde{N}(y, \rho)}{\widetilde{D}(y, \rho)}\right)=-\frac{\left(\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)\right) \widetilde{D}(y, \rho)-\widetilde{N}(y, \rho)\left(\frac{\partial}{\partial \rho} \widetilde{D}(y, \rho)\right)}{\widetilde{D}^{2}(y, \rho)}
$$

or using (23),

$$
\frac{\partial}{\partial \rho}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\widetilde{N}(y, \rho)}{\widetilde{D}(y, \rho)}\right)=-\frac{\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)-\frac{\lambda_{1}}{\lambda_{2}} y \frac{\partial}{\partial \rho} \widetilde{D}(y, \rho)}{\widetilde{D}(y, \rho)} .
$$

Analogously,

$$
\frac{\partial}{\partial y}\left(\frac{\lambda_{1}}{\lambda_{2}} y-\frac{\widetilde{N}(y, \rho)}{\widetilde{D}(y, \rho)}\right)=\frac{\lambda_{1}}{\lambda_{2}}-\frac{\left(\frac{\partial}{\partial y} \widetilde{N}(y, \rho)\right)-\frac{\lambda_{1}}{\lambda_{2}} y\left(\frac{\partial}{\partial y} \widetilde{D}(y, \rho)\right)}{\widetilde{D}(y, \rho)}
$$

Therefore, (27) is equivalent to

$$
\frac{\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)-\frac{\lambda_{1}}{\lambda_{2}} y \frac{\partial}{\partial \rho} \widetilde{D}(y, \rho)}{\widetilde{D}(y, \rho)}>-\frac{\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)}{\frac{\partial}{\partial y} \widetilde{N}(y, \rho)}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{\left(\frac{\partial}{\partial y} \widetilde{N}(y, \rho)\right)-\frac{\lambda_{1}}{\lambda_{2}} y\left(\frac{\partial}{\partial y} \widetilde{D}(y, \rho)\right)}{\widetilde{D}(y, \rho)}\right)
$$

or,

$$
\begin{equation*}
-\frac{y \frac{\partial}{\partial \rho} \widetilde{D}(y, \rho)}{\widetilde{D}(y, \rho)}>-\frac{\frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)}{\frac{\partial}{\partial y} \widetilde{N}(y, \rho)}\left(1+\frac{y \frac{\partial}{\partial y} \widetilde{D}(y, \rho)}{\widetilde{D}(y, \rho)}\right) . \tag{28}
\end{equation*}
$$

Moreover, recall that $a_{2}=\widetilde{D}(y, \rho) / \lambda_{2}$. The positiveness of $a_{2}$ tells us that $\widetilde{D}(y, \rho)>0$. After some algebra, we have that $\frac{\partial}{\partial \rho} \widetilde{D}(y, \rho)<0, \frac{\partial}{\partial \rho} \widetilde{N}(y, \rho)<0$ and $\frac{\partial}{\partial y} \widetilde{D}(y, \rho)>0$. Hence, we conclude that the LHS of (28) is positive, whereas the $R H S$ of (28) is negative since $\frac{\partial}{\partial y} \tilde{N}(y, \rho)<0$. Consequently, the fact that (28) is satisfied allows us to conclude that $\frac{\partial a_{1}}{\partial \rho}<0$.

Concerning the effect of $\rho$ on $c_{1}$, recall that $c_{1}=a_{1} / \Delta_{1}=\left(1+\rho+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right) a_{1} /(1+\rho)$. This expression tells us that $c_{1}$ is the product of two decreasing positive functions in $\rho$. Therefore, $\frac{\partial c_{1}}{\partial \rho}<0$.

Next, we prove that $\frac{\partial}{\partial \sigma_{\varepsilon_{1}}^{2}}\left(d_{1} / d_{2}\right)<0$ and $\frac{\partial}{\partial \sigma_{\varepsilon_{2}}^{2}}\left(d_{1} / d_{2}\right)>0$. From the expressions of $d_{1}$, $d_{2}$, and $z$, it follows that $d_{1} / d_{2}=\left(n_{1} z+\left(n_{2}-1\right)\right) /\left(\left(n_{1}-1\right) z+n_{2}\right)$. Applying the chain rule, we get $\frac{\partial}{\partial \sigma_{\varepsilon_{i}}^{2}}\left(d_{1} / d_{2}\right)=\frac{\partial}{\partial z}\left(d_{1} / d_{2}\right) \frac{\partial z}{\partial \sigma_{\varepsilon_{i}}^{2}}$. As $\frac{\partial}{\partial z}\left(d_{1} / d_{2}\right)>0$, we know that the sign of $\frac{\partial}{\partial \sigma_{\varepsilon_{i}}^{2}}\left(d_{1} / d_{2}\right)$ is the same as the sign of $\frac{\partial z}{\partial \sigma_{\varepsilon_{i}}^{2}}$. Applying the Implicit Function Theorem, $\frac{\partial z}{\partial \sigma \sigma_{i}^{2}}=$ $-\frac{\partial}{\partial \sigma_{\varepsilon_{i}}^{2}}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right) / \frac{\partial}{\partial z}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)$. Direct computation yields that $\frac{\partial}{\partial \sigma_{\varepsilon_{1}}^{2}}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)>0$ and $\frac{\partial}{\partial \sigma_{\varepsilon_{2}}^{2}}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)<0$. In addition, since $\frac{\partial}{\partial z}\left(\frac{\lambda_{1}}{\lambda_{2}}-\frac{N(z)}{D(z)}\right)>0$, we obtain $\frac{\partial z}{\partial \sigma_{\varepsilon_{1}}^{2}}<0$ and $\frac{\partial z}{\partial \sigma_{\varepsilon_{2}}^{2}}>0$, and hence, we conclude that $\frac{\partial}{\partial \sigma_{\varepsilon_{1}}^{2}}\left(d_{1} / d_{2}\right)<0$ and $\frac{\partial}{\partial \sigma_{\varepsilon_{2}}^{2}}\left(d_{1} / d_{2}\right)>0$. Analogously, the negativeness of $\frac{\partial z}{\partial \lambda_{1}}$ and the positiveness of $\frac{\partial z}{\partial \lambda_{2}}$ allows us to conclude that $\frac{\partial}{\partial \lambda_{1}}\left(d_{1} / d_{2}\right)<0$ and $\frac{\partial}{\partial \lambda_{2}}\left(d_{1} / d_{2}\right)>0$.

Now, we study how $a_{1}$ and $c_{1}$ vary with a change in $\sigma_{\varepsilon_{i}}^{2}, i=1,2$. In order to do that first we analyze the effect of $\sigma_{\varepsilon_{i}}^{2}$ on $d_{1}$ and $d_{2}$. From Proposition 1, we know that $d_{i}=$ $\left(\left(n_{i}-1\right) c_{i}+n_{j} c_{j}\right)^{-1}$ and $a_{i}=\Delta_{i} c_{i}>0, i=1,2$. Therefore, $d_{i}=\left(\left(n_{i}-1\right) \Delta_{i}^{-1} a_{i}+n_{j} \Delta_{j}^{-1} a_{j}\right)^{-1}$, $i, j=1,2, j \neq i$. Substituting the expressions of (12) and the expression for $\Delta_{i}$ given in Lemma A2, it follows that

$$
d_{i}=\left(\frac{\left(n_{i}-1\right) n_{j}}{n_{j} \Upsilon_{i}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(\Upsilon_{i}-1\right)\left(d_{j}+\lambda_{j}\right)}+\frac{n_{j} n_{i}}{n_{i} \Upsilon_{j}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(\Upsilon_{j}-1\right)\left(d_{i}+\lambda_{i}\right)}\right)^{-1}
$$

where $\Upsilon_{i}=\Xi_{j} /\left(\Xi_{j}-\Psi_{i}\right)=\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right) /\left((1-\rho)\left(1+\rho+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)\right)>1, i, j=1,2, j \neq i$. Therefore, we derive the following equations that are satisfied in equilibrium: $F_{i}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)=$ $0, i=1,2$, where
$F_{i}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)=\frac{\left(n_{i}-1\right) n_{j} d_{i}}{n_{j} \Upsilon_{i}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(\Upsilon_{i}-1\right)\left(d_{j}+\lambda_{j}\right)}+\frac{n_{i} n_{j} d_{i}}{n_{i} \Upsilon_{j}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(\Upsilon_{j}-1\right)\left(d_{i}+\lambda_{i}\right)}-1$,
$i, j=1,2, j \neq i$. Let $D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)$ denote the following matrix:

$$
\left(\begin{array}{ll}
\frac{\partial F_{1}}{\partial d_{1}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right) & \frac{\partial F_{1}}{\partial d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right) \\
\frac{\partial F_{2}}{\partial d_{1}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right) & \frac{\partial F_{2}}{\partial d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)
\end{array}\right) .
$$

Direct computation yields

$$
\begin{aligned}
& \quad \frac{\partial F_{i}}{\partial d_{i}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)= \\
& \frac{n_{j}\left(n_{i}-1\right)\left(n_{i}\left(\lambda_{j}+d_{j}\right)\left(\Upsilon_{i}-1\right)+\Upsilon_{i} \lambda_{i} n_{j}\right)}{\left(n_{j} \Upsilon_{i}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(\Upsilon_{i}-1\right)\left(d_{j}+\lambda_{j}\right)\right)^{2}}+\frac{n_{i} n_{j}\left(\lambda_{i} n_{j}\left(\Upsilon_{j}-1\right)+\Upsilon_{j} n_{i}\left(\lambda_{j}+d_{j}\right)\right)}{\left(n_{i} \Upsilon_{j}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(\Upsilon_{j}-1\right)\left(d_{i}+\lambda_{i}\right)\right)^{2}}, \text { and }
\end{aligned}
$$

$$
\frac{\partial F_{i}}{\partial d_{j}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)=
$$

$$
-\left(\frac{d_{i} n_{i} n_{j}\left(n_{i}-1\right)\left(\Upsilon_{i}-1\right)}{\left(n_{j} \Upsilon_{i}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(\Upsilon_{i}-1\right)\left(d_{j}+\lambda_{j}\right)\right)^{2}}+\frac{d_{i} n_{i}^{2} n_{j} \Upsilon_{j}}{\left(n_{i} \Upsilon_{j}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(\Upsilon_{j}-1\right)\left(d_{i}+\lambda_{i}\right)\right)^{2}}\right)
$$

$i, j=1,2, j \neq i$. After some tedious algebra, the determinant of $D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)$ is given by

$$
\begin{gathered}
n_{1} n_{2}^{2}\left(n_{1}-1\right) \frac{\lambda_{1} \Upsilon_{1} n_{2}+\lambda_{2} n_{1}\left(\Upsilon_{1}-1\right)}{\left(n_{2} \Upsilon_{1}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(\Upsilon_{1}-1\right)\left(d_{2}+\lambda_{2}\right)\right)^{3}}+ \\
n_{1}^{2} n_{2}\left(n_{2}-1\right) \frac{\lambda_{1} n_{2}\left(\Upsilon_{2}-1\right)+\lambda_{2} \Upsilon_{2} n_{1}}{\left(n_{1} \Upsilon_{2}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(\Upsilon_{2}-1\right)\left(d_{1}+\lambda_{1}\right)\right)^{3}}+ \\
\frac{n_{1} n_{2}}{\left(n_{2} \Upsilon_{1}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(\Upsilon_{1}-1\right)\left(d_{2}+\lambda_{2}\right)\right)^{2}\left(n_{1} \Upsilon_{2}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(\Upsilon_{2}-1\right)\left(d_{1}+\lambda_{1}\right)\right)^{2}} \times \\
\left(\left(n_{1} n_{2}+\left(n_{1}-1\right)\left(n_{2}-1\right)\right)\left(\lambda_{1} n_{2}\left(\Upsilon_{2}-1\right)+\lambda_{2} \Upsilon_{2} n_{1}\right)\left(\lambda_{1} \Upsilon_{1} n_{2}+\lambda_{2} n_{1}\left(\Upsilon_{1}-1\right)\right)+\right. \\
n_{2}\left(d_{2} n_{1}\left(\Upsilon_{1}-1\right)+\Upsilon_{1} d_{1} n_{2}\right)\left(\lambda_{1}\left(\Upsilon_{2}-1\right)\left(n_{1}-1\right)\left(n_{2}-1\right)+\lambda_{2} \Upsilon_{2} n_{1}^{2}\right)+ \\
\left.n_{1}\left(d_{1} n_{2}\left(\Upsilon_{2}-1\right)+\Upsilon_{2} d_{2} n_{1}\right)\left(\lambda_{1} \Upsilon_{1} n_{2}^{2}+\lambda_{2}\left(\Upsilon_{1}-1\right)\left(n_{1}-1\right)\left(n_{2}-1\right)\right)\right) .
\end{gathered}
$$

Notice that all the summands in the previous expression are positive. In particular, the determinant of $D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)$ is not null and, therefore, this matrix is invertible. Hence, we can apply the Implicit Function Theorem, we have

$$
\left(\begin{array}{cc}
\frac{\partial d_{1}}{\partial \sigma_{\varepsilon_{1}}^{2}} & \frac{\partial d_{1}}{\partial \sigma_{\varepsilon_{2}}^{2}} \\
\frac{\partial d_{2}}{\partial \sigma_{\varepsilon_{1}}^{2}} & \frac{\partial d_{2}}{\partial \sigma_{\varepsilon_{1}}^{2}}
\end{array}\right)=
$$

$$
-\left(D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)\right)^{-1}\left(\begin{array}{cc}
\frac{\partial F_{1}}{\partial \sigma_{\varepsilon_{1}}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right) & \frac{\partial F_{1}}{\partial \sigma_{\varepsilon_{2}}^{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)  \tag{29}\\
\frac{\partial F_{2}}{\partial \sigma_{\varepsilon_{1}}^{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right) & \frac{\partial F_{2}}{\partial \sigma_{\varepsilon_{2}}^{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)
\end{array}\right) .
$$

Notice that

$$
\left(D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)\right)^{-1}=\left(\begin{array}{cc}
\frac{\frac{\partial F_{2}}{\partial d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)}{\operatorname{det}\left(D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)\right)} & -\frac{\frac{\partial F_{1}}{\partial d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)}{\operatorname{det}\left(D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)\right)} \\
-\frac{\frac{\partial F_{2}}{\partial d_{1}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)}}{\operatorname{det}\left(D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)\right)} & \frac{\frac{\partial F_{1}}{\partial d_{1}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)}{\operatorname{det}\left(D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)\right)}
\end{array}\right)
$$

Hence, we know that all the elements of $\left(D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)\right)^{-1}$ are positive. Moreover,

$$
\begin{aligned}
\frac{\partial F_{i}}{\partial \sigma_{\varepsilon_{i}}^{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right) & =-\frac{d_{i} n_{j}\left(n_{i}-1\right)\left(n_{j}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(d_{j}+\lambda_{j}\right)\right)}{\left(n_{j} \Upsilon_{i}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(\Upsilon_{i}-1\right)\left(d_{j}+\lambda_{j}\right)\right)^{2}} \frac{\rho(1+\rho)}{(1-\rho) \sigma_{\theta}^{2}\left(1+\rho+\widehat{\sigma}_{\varepsilon_{i}}^{2}\right)^{2}}, \text { and } \\
\frac{\partial F_{i}}{\partial \sigma_{\varepsilon_{j}}^{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right) & =-\frac{d_{i} n_{i} n_{j}\left(n_{j}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(d_{j}+\lambda_{j}\right)\right)}{\left(n_{i} \Upsilon_{j}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(\Upsilon_{j}-1\right)\left(d_{i}+\lambda_{i}\right)\right)^{2}} \frac{\rho(1+\rho)}{(1-\rho) \sigma_{\theta}^{2}\left(1+\rho+\widehat{\sigma}_{\varepsilon_{j}}^{2}\right)^{2}}
\end{aligned}
$$

$i, j=1,2, j \neq i$. Using all these expressions, (29) implies that $\frac{\partial d_{i}}{\partial \sigma_{\varepsilon_{i}}^{2}}>0$ and $\frac{\partial d_{i}}{\partial \sigma_{\varepsilon_{j}}^{2}}>0$.
Next, we study the comparative statics of $c_{1}$ and $c_{2}$ with respect to $\sigma_{\varepsilon_{1}}^{2}$. Recall that

$$
\begin{aligned}
& c_{1}=a_{1} / \Delta_{1}=n_{2} /\left(n_{2} \Upsilon_{1}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(\Upsilon_{1}-1\right)\left(d_{2}+\lambda_{2}\right)\right) \text { and } \\
& c_{2}=a_{2} / \Delta_{2}=n_{1} /\left(n_{1} \Upsilon_{2}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(\Upsilon_{2}-1\right)\left(d_{1}+\lambda_{1}\right)\right) .
\end{aligned}
$$

Using the fact that $\Upsilon_{1}, d_{1}$, and $d_{2}$ are increasing in $\sigma_{\varepsilon_{1}}^{2}$ and that $\Upsilon_{2}$ is independent of $\sigma_{\varepsilon_{1}}^{2}$, we have the denominators of the previous expressions are increasing in $\sigma_{\varepsilon_{1}}^{2}$, which allows us to conclude that $c_{1}$ and $c_{2}$ are decreasing in $\sigma_{\varepsilon_{1}}^{2}$. Combining these results with the fact that $\Delta_{1}$ is decreasing in $\sigma_{\varepsilon_{1}}^{2}$ and $\Delta_{2}$ is independent of $\sigma_{\varepsilon_{1}}^{2}$, it follows that $a_{1}$ and $a_{2}$ are decreasing in $\sigma_{\varepsilon_{1}}^{2}$, since $a_{1}=\Delta_{1} c_{1}$ and $a_{2}=\Delta_{2} c_{2}$.

Finally, concerning h), notice that doing a similar reasoning as before we derive the following equations that are satisfied in equilibrium: $F_{i}\left(n_{1}, n_{2}, d_{1}, d_{2}\right)=0, i=1,2$, where $F_{i}\left(n_{1}, n_{2}, d_{1}, d_{2}\right)=F_{i}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right), i, j=1,2, j \neq i$. Hence,

$$
\left(\begin{array}{ll}
\frac{\partial d_{1}}{\partial n_{1}} & \frac{\partial d_{1}}{\partial n_{2}} \\
\frac{\partial d_{2}}{\partial n_{1}} & \frac{\partial d_{2}}{\partial n_{2}}
\end{array}\right)=-\left(D F_{d_{1}, d_{2}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right)\right)^{-1}\left(\begin{array}{ll}
\frac{\partial F_{1}}{\partial n_{1}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right) & \frac{\partial F_{1}}{\partial n_{2}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right) \\
\frac{\partial F_{2}}{\partial n_{1}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right) & \frac{\partial F_{2}}{\partial n_{2}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right)
\end{array}\right) .
$$

In addition,

$$
\begin{aligned}
& \quad \frac{\partial F_{i}}{\partial n_{i}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right)= \\
& d_{i} n_{j} \frac{\left(\Upsilon_{i}-1\right)\left(d_{j}+\lambda_{j}\right)+\Upsilon_{i} n_{j}\left(d_{i}+\lambda_{i}\right)}{\left(n_{j} \Upsilon_{i}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(\Upsilon_{i}-1\right)\left(d_{j}+\lambda_{j}\right)\right)^{2}}+\frac{d_{i} n_{j}^{2}\left(\Upsilon_{j}-1\right)\left(d_{i}+\lambda_{i}\right)}{\left(n_{i} \Upsilon_{j}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(\Upsilon_{j}-1\right)\left(d_{i}+\lambda_{i}\right)\right)^{2}}, \text { and }
\end{aligned}
$$

$$
\begin{aligned}
& \frac{\partial F_{i}}{\partial n_{j}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right)= \\
& \quad \frac{d_{i} n_{i}\left(\Upsilon_{i}-1\right)\left(n_{i}-1\right)\left(d_{j}+\lambda_{j}\right)}{\left(n_{j} \Upsilon_{i}\left(d_{i}+\lambda_{i}\right)+n_{i}\left(\Upsilon_{i}-1\right)\left(d_{j}+\lambda_{j}\right)\right)^{2}}+\frac{\Upsilon_{j} d_{i} n_{i}^{2}\left(d_{j}+\lambda_{j}\right)}{\left(n_{i} \Upsilon_{j}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(\Upsilon_{j}-1\right)\left(d_{i}+\lambda_{i}\right)\right)^{2}},
\end{aligned}
$$

$i, j=1,2, j \neq i$. Therefore, all these partial derivatives are positive. Combining this result with the positiveness of all the elements of $\left(D F_{d_{1}, d_{2}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right)\right)^{-1},{ }^{22}$ we conclude that $\frac{\partial d_{i}}{\partial n_{i}}<0$ and $\frac{\partial d_{i}}{\partial n_{j}}<0, i, j=1,2, j \neq i$.

Proof of Corollary 2: (i) Suppose that $\sigma_{\varepsilon_{1}}^{2} \geq \sigma_{\varepsilon_{2}}^{2}, \lambda_{1} \geq \lambda_{2}$, and $n_{1} \geq n_{2}$. Using the expressions of $\Xi_{i}$ and $\Delta_{i}, i=1,2$, it is easy to see that in this case $\Xi_{2} \Delta_{2}^{-1}>\Xi_{1} \Delta_{1}^{-1}$. Next, we distinguish two cases:

Case 1: $\left(n_{1}+n_{2}-2\right) n_{1} /\left(\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)\right) \geq 1-\Xi_{2} \Delta_{2}^{-1}$. Evaluating the polynomial $p(z)$, stated in the proof of Lemma A2, at $z=1$, we have that
$p(1)=\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)^{2} \times$
$\left(n_{2}\left(\frac{\left(n_{1}+n_{2}-2\right) n_{1}}{\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)}-\left(1-\Xi_{2} \Delta_{2}^{-1}\right)\right) \lambda_{1}-n_{1}\left(\frac{\left(n_{1}+n_{2}-2\right) n_{2}}{\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)}-\left(1-\Xi_{1} \Delta_{1}^{-1}\right)\right) \lambda_{2}\right)$.
As $\left(n_{1}+n_{2}-2\right) n_{1} /\left(\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)\right) \geq 1-\Xi_{2} \Delta_{2}^{-1}$ and $\lambda_{1} \geq \lambda_{2}$,

$$
p(1) \geq\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)^{2}\left(\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1}-n_{2}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)\right) \lambda_{2}
$$

and since $n_{1} \geq n_{2}, p(1) \geq\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)^{2} \lambda_{2} n_{2}\left(\Xi_{2} \Delta_{2}^{-1}-\Xi_{1} \Delta_{1}^{-1}\right) \geq 0$. This implies that $z \leq 1$, and therefore, $c_{1} \leq c_{2}$. In addition, using the expressions of $d_{1}$ and $d_{2}$, we get $\operatorname{sgn}\left\{d_{1}-d_{2}\right\}=\operatorname{sgn}\left\{c_{1}-c_{2}\right\}$, which implies $d_{1} \leq d_{2}$. Finally, notice that $\Delta_{1} \leq \Delta_{2}$ whenever $\sigma_{\varepsilon_{1}}^{2} \geq \sigma_{\varepsilon_{2}}^{2}$. Hence, $a_{1} / a_{2}=z \Delta_{1} / \Delta_{2} \leq 1$.
Case 2: $\left(n_{1}+n_{2}-2\right) n_{1} /\left(\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)\right)<1-\Xi_{2} \Delta_{2}^{-1}$. Notice that

$$
\frac{\left(n_{1}+n_{2}-2\right) n_{2}}{\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)}-\left(1-\Xi_{1} \Delta_{1}^{-1}\right) \leq \frac{\left(n_{1}+n_{2}-2\right) n_{1}}{\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)}-\left(1-\Xi_{2} \Delta_{2}^{-1}\right)
$$

since $\Xi_{2} \Delta_{2}^{-1}>\Xi_{1} \Delta_{1}^{-1}$ and $n_{1} \geq n_{2}$. Thus, in this case we have that

$$
\begin{aligned}
& q_{N}(1)=\left(n_{1}+n_{2}-1\right)\left(n_{1}+n_{2}\right)\left(\frac{\left(n_{1}+n_{2}-2\right) n_{2}}{\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)}-\left(1-\Xi_{1} \Delta_{1}^{-1}\right)\right)<0 \text { and } \\
& q_{D}(1)=\left(n_{1}+n_{2}-1\right)\left(n_{1}+n_{2}\right)\left(\frac{\left(n_{1}+n_{2}-2\right) n_{1}}{\left(n_{1}+n_{2}\right)\left(n_{1}+n_{2}-1\right)}-\left(1-\Xi_{2} \Delta_{2}^{-1}\right)\right)<0
\end{aligned}
$$

Taking into account the shape of these polynomials, the previous two inequalities imply that $\bar{z}_{D}>1>\bar{z}_{N}$. However, Proposition 2 indicates that in this case there is no equilibrium.
${ }^{22}$ Notice that $\left(D F_{d_{1}, d_{2}}\left(n_{1}, n_{2}, d_{1}, d_{2}\right)\right)^{-1}=\left(D F_{d_{1}, d_{2}}\left(\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}, d_{1}, d_{2}\right)\right)^{-1}$.
(ii) By virtue of (17), the inequality $d_{1}+\lambda_{1}>d_{2}+\lambda_{2}$ is equivalent to

$$
\left(\Xi_{1}-\frac{n_{1}}{n_{2}} \frac{a_{1}}{a_{2}} \Psi_{1}\right) / a_{1}>\left(\Xi_{2}-\frac{n_{2}}{n_{1}} \frac{a_{2}}{a_{1}} \Psi_{2}\right) / a_{2}
$$

Using (14) and (20), and after some algebra, the previous inequality is equivalent to

$$
\begin{equation*}
z<\frac{\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)}{\Xi_{2} \Delta_{2}^{-1}+n_{1} n_{2}^{-1}\left(1-\Xi_{1} \Delta_{1}^{-1}\right)} \tag{30}
\end{equation*}
$$

We distinguish two cases:
Case I: $\frac{\Xi_{2} \Delta_{2}^{-1}+\Xi_{1} \Delta_{1}^{-1}-1}{\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} n_{2}^{-1}+\Xi_{2} \Delta_{2}^{-1}} \leq \frac{\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)}{n_{1}\left(\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)\right)+\left(n_{2}-1\right)\left(\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} n_{2}^{-1}+\Xi_{2} \Delta_{2}^{-1}\right)}$. Using the expressions of $\Xi_{i}$ and $\Delta_{i}$, we get that the previous inequality is equivalent to

$$
\frac{(1-\rho) n_{1} n_{2}\left(1+\rho+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)}{n_{2}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)+n_{1} \rho \widehat{\sigma}_{\varepsilon_{1}}^{2}}+\frac{(1-\rho) n_{1}\left(n_{2}-1\right)\left(1+\rho+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)}{n_{1}\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)+n_{2} \rho \widehat{\sigma}_{\varepsilon_{2}}^{2}} \leq 1
$$

Moreover, after some algebra, we have that $q_{D}\left(\frac{\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)}{\Xi_{2} \Delta_{2}^{-1}+n_{1} n_{2}^{-1}\left(1-\Xi_{1} \Delta_{1}^{-1}\right)}\right) \leq 0$. Consequently, $\frac{\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)}{\Xi_{2} \Delta_{2}^{-1}+n_{1} n_{2}^{-1}\left(1-\Xi_{1} \Delta_{1}^{-1}\right)}<\bar{z}_{D}<z$, which implies that in this case $d_{1}+\lambda_{1}<d_{2}+\lambda_{2}$ holds.

Case II: $\frac{\Xi_{2} \Delta_{2}^{-1}+\Xi_{1} \Delta_{1}^{-1}-1}{\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} n_{2}^{-1}+\Xi_{2} \Delta_{2}^{-1}}>\frac{\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)}{n_{1}\left(\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)\right)+\left(n_{2}-1\right)\left(\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} n_{2}^{-1}+\Xi_{2} \Delta_{2}^{-1}\right)}$. In this case, taking into account that $z$ is the unique positive solution of $(21),(30)$ is equivalent to

$$
\frac{\lambda_{1}}{\lambda_{2}}>\frac{\frac{\Xi_{2} \Delta_{2}^{-1}+\Xi_{1} \Delta_{1}^{-1}-1}{\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} n_{2}^{-1}+\Xi_{2} \Delta_{2}^{-1}}-\frac{\left(\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)\right)}{\left(n_{1}-1\right)\left(\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)\right)+n_{2}\left(\left(1-\Xi_{1} \Delta_{1}^{-1}\right) \frac{n_{1}}{n_{2}}+\Xi_{2} \Delta_{2}^{-1}\right)}}{\frac{\Xi_{2} \Delta_{2}^{-1}+\Xi_{1} \Delta_{1}^{-1}-1}{\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} n_{2}^{-1}+\Xi_{2}^{-1}\left(1-\Xi_{2}^{-1} \Delta_{2}^{-1}\right)}-\frac{\Xi_{1}}{n_{1}\left(\Xi_{1} \Delta_{1}^{-1}+n_{2} n_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right)\right)+\left(n_{2}-1\right)\left(\left(1-\Xi_{1} \Delta_{1}^{-1}\right) n_{1} n_{2}^{-1}+\Xi_{2} \Delta_{2}^{-1}\right)}} .
$$

Taking into account that $\sigma_{\varepsilon_{1}}^{2} \geq \sigma_{\varepsilon_{2}}^{2}, \lambda_{1} \geq \lambda_{2}$, and $n_{1} \geq n_{2}$, and after some algebra, we get that both sides of the inequality are higher than (or equal to) 1 . Therefore, if the value of $\lambda_{1} / \lambda_{2}$ is high enough, then we obtain $d_{1}+\lambda_{1}>d_{2}+\lambda_{2}$. Otherwise, the opposite inequality holds.

Proof of Proposition 4: (i) First, suppose that $\bar{\theta}_{1}=\bar{\theta}_{2}$. Then $E[p]=\bar{\theta}_{1}-Q /\left(\frac{n_{1}}{d_{1}+\lambda_{1}}+\frac{n_{2}}{d_{2}+\lambda_{2}}\right)$. From Proposition 3 we know that $d_{i}$ and $d_{j}$ decrease with $n_{i}$, and increase with $\sigma_{\varepsilon_{i}}^{2}, \lambda_{i}$, and $\rho$. Using these results in the previous expression, we conclude that the expected price is increasing in $n_{i}$ and decreasing in $\lambda_{i}, \sigma_{\varepsilon_{i}}^{2}$, and $\rho$.

Now, suppose that $\bar{\theta}_{1} \neq \bar{\theta}_{2}$. The results we have just derived may not hold if $\left|\bar{\theta}_{1}-\bar{\theta}_{2}\right|$ is large enough. For example, let us focus on the relationship between the expected price and
$n_{1}$. To study this relationship, we first show that $n_{2}\left(d_{1}+\lambda_{1}\right) /\left(n_{1}\left(d_{2}+\lambda_{2}\right)\right)$ decreases with $n_{1}$. Recall that

$$
d_{2}=1 /\left(\frac{n_{1} n_{2}}{n_{2} \Upsilon_{1}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(\Upsilon_{1}-1\right)\left(d_{2}+\lambda_{2}\right)}+\frac{\left(n_{2}-1\right) n_{1}}{n_{1} \Upsilon_{2}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(\Upsilon_{2}-1\right)\left(d_{1}+\lambda_{1}\right)}\right)
$$

Hence,

$$
1=\left(\frac{n_{2}}{\Upsilon_{1} \frac{n_{2}\left(d_{1}+\lambda_{1}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)}+\Upsilon_{1}-1}+\frac{n_{2}-1}{\Upsilon_{2}+\left(\Upsilon_{2}-1\right) \frac{n_{2}\left(d_{1}+\lambda_{1}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)}}\right)^{-1}+\frac{\lambda_{2}}{d_{2}+\lambda_{2}}
$$

The fact that $d_{2}$ decreases with $n_{1}$ implies that $\lambda_{2} /\left(d_{2}+\lambda_{2}\right)$ increases with $n_{1}$. Then, the previous inequality tells us that $\frac{n_{2}}{\Upsilon_{1} \frac{n_{2}\left(d_{1}+\lambda_{1}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)}+\Upsilon_{1}-1}+\frac{n_{2}-1}{\Upsilon_{2}+\left(\Upsilon_{2}-1 \frac{n_{2}\left(d_{1}+\lambda_{1}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)}\right.}$ increases with $n_{1}$. For this to be possible, $\frac{n_{2}\left(d_{1}+\lambda_{1}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)}$ needs to be decreasing in $n_{1}$. Given that the expected price satisfies
$E[p]=\left(1+\frac{n_{2}\left(d_{1}+\lambda_{1}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)}\right)^{-1} \bar{\theta}_{1}+\left(1-\left(1+\frac{n_{2}\left(d_{1}+\lambda_{1}\right)}{n_{1}\left(d_{2}+\lambda_{2}\right)}\right)^{-1}\right) \bar{\theta}_{2}-\left(\frac{n_{1}}{d_{1}+\lambda_{1}}+\frac{n_{2}}{d_{2}+\lambda_{2}}\right)^{-1} Q$,
we have that the relationship between the expected price and $n_{1}$ is ambiguous. For instance, if $\bar{\theta}_{2}$ is low enough, then the fact that $d_{1}, d_{2}$, and $n_{2}\left(d_{1}+\lambda_{1}\right) /\left(n_{1}\left(d_{2}+\lambda_{2}\right)\right)$ are decreasing in $n_{1}$ allows us to conclude that the expected price increases with $n_{1}$. However, if $\bar{\theta}_{2}$ is large and $\bar{\theta}_{1}$ and $Q$ are low enough, then the expected price decreases with $n_{1}$.
(ii) From the expression for the expected revenue it follows that $Q E[p]$ increases with $\bar{\theta}_{i}$, $i=1,2$, and with $Q$, whenever $Q<\frac{n_{1}}{d_{1}+\lambda_{1}} \bar{\theta}_{1}+\frac{n_{2}}{d_{2}+\lambda_{2}} \bar{\theta}_{2}$, or equivalently, $E[p]>0$. In addition, using the expression for the expected price, it follows that

$$
\left(\min \left\{\bar{\theta}_{1}, \bar{\theta}_{2}\right\}-\frac{Q}{\frac{n_{1}}{d_{1}+\lambda_{1}}+\frac{n_{2}}{d_{2}+\lambda_{2}}}\right) \leq E[p] \leq\left(\max \left\{\bar{\theta}_{1}, \bar{\theta}_{2}\right\}-\frac{Q}{\frac{n_{1}}{d_{1}+\lambda_{1}}+\frac{n_{2}}{d_{2}+\lambda_{2}}}\right)
$$

Notice that left-hand side $(L H S)$ and the right-hand side $(R H S)$ of this expression corresponds to the expected revenue in an auction where all participants have an expected valuation of $\min \left\{\bar{\theta}_{1}, \bar{\theta}_{2}\right\}$ and of max $\left\{\bar{\theta}_{1}, \bar{\theta}_{2}\right\}$, respectively. Using Proposition 4(i), we know that both LHS and RHS increase with $n_{i}$ but decrease with $\lambda_{i}$ and $\sigma_{\varepsilon_{i}}^{2}$. Hence, we obtain that $Q E[p]$ is lower than the expected revenue of the symmetric auction in which both groups are ex-ante identical, with large size (each group with $\max \left\{n_{1}, n_{2}\right\}$ bidders), with high expected valuation $\left(\max \left\{\bar{\theta}_{1}, \bar{\theta}_{2}\right\}\right)$, low transaction costs $\left(\min \left\{\lambda_{1}, \lambda_{2}\right\}\right)$, and precise signals $\left(\min \left\{\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}\right\}\right)$, and larger than the expected revenue of the symmetric auction in which both groups are ex-ante identical but with the opposite characteristics (i.e., $\min \left\{n_{1}, n_{2}\right\}, \min \left\{\bar{\theta}_{1}, \bar{\theta}_{2}\right\}, \max \left\{\lambda_{1}, \lambda_{2}\right\}$, and $\left.\max \left\{\sigma_{\varepsilon_{1}}^{2}, \sigma_{\varepsilon_{2}}^{2}\right\}\right)$.

Proof of Proposition 5: Direct computation yields
$\bar{z}_{N}=\frac{n_{2}\left(\left(n_{1}-1\right)\left(2 \Xi_{1} \Delta_{1}^{-1}-1\right)-\left(2-\Xi_{1} \Delta_{1}^{-1}\right)+\sqrt{\left(2-\Xi_{1} \Delta_{1}^{-1}\right)^{2}+\left(n_{1}-1\right)\left(n_{1}+3-6 \Xi_{1} \Delta_{1}^{-1}\right)}\right)}{2 n_{1}\left(n_{1}-1\right)\left(1-\Xi_{1} \Delta_{1}^{-1}\right)}$
and

$$
\bar{z}_{D}=\frac{n_{2}+1-\Xi_{2} \Delta_{2}^{-1}\left(2 n_{2}-1\right)+\sqrt{\left(2-\Xi_{2} \Delta_{2}^{-1}\right)^{2}+\left(n_{2}-1\right)\left(n_{2}+3-6 \Xi_{2} \Delta_{2}^{-1}\right)}}{2 \Xi_{2} \Delta_{2}^{-1} n_{1}}
$$

Thus, when $n_{1}$ tends to infinity, note that $\lim _{n_{1} \rightarrow \infty} \bar{z}_{N}=\lim _{n_{1} \rightarrow \infty} \bar{z}_{D}=0$. Furthermore, using the previous expressions and after some algebra, the necessary and sufficient condition for the existence of an equilibrium (i.e., $\lim _{n_{1} \rightarrow \infty} \frac{\bar{z}_{N}}{\bar{z}_{D}}>1$ ) is equivalent to $n_{2}>\bar{n}_{2}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)$, where

$$
\bar{n}_{2}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)=\frac{\left((2-\rho) \widehat{\sigma}_{\varepsilon_{2}}^{2}+2\left(1-\rho^{2}\right)\right) \widehat{\sigma}_{\varepsilon_{1}}^{2} \rho}{\left(1-\rho^{2}\right)\left(\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}\right)}
$$

Moreover, taking the limit in (21), it follows that $\lim _{n_{1} \rightarrow \infty} z=0$ and

$$
\begin{equation*}
\lim _{n_{1} \rightarrow \infty} n_{1} z=n_{2} \Xi_{1} \Delta_{1}^{-1} /\left(1-\Xi_{1} \Delta_{1}^{-1}\right) \tag{31}
\end{equation*}
$$

Using the expressions included in the statement of Lemma A2, and after some tedious algebra, we get $\lim _{n_{1} \rightarrow \infty} b_{1}=q, \lim _{n_{1} \rightarrow \infty} a_{1}=0, \lim _{n_{1} \rightarrow \infty} c_{1}=0, \lim _{n_{1} \rightarrow \infty} a_{2}=\Delta_{2} \lim _{n_{1} \rightarrow \infty} c_{2}$,

$$
\begin{align*}
\lim _{n_{1} \rightarrow \infty} b_{2}= & \frac{\widehat{\sigma}_{\varepsilon_{2}}^{2}\left(\frac{\left(n_{2}-1\right)\left(1-\rho^{2}\right)}{\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)}+\frac{\left(1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{1}}^{2}(1-2 \rho)\right)}{\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}}\right)\left(q \rho \lambda_{1}+\bar{\theta}_{2}-\rho \bar{\theta}_{1}\right)}{(1-\rho) \lambda_{2}\left(n_{2}(1+\rho)-\frac{\widehat{\sigma}_{\varepsilon_{1}}^{2} \rho\left(1+\rho+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)}{\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}}\right)}+ \\
& +q \frac{\rho^{2} \widehat{\sigma}_{\varepsilon_{2}}^{2} \widehat{\sigma}_{\varepsilon_{1}}^{2}}{n_{2}\left(1-\rho^{2}\right)\left(\left(1+\widehat{\sigma}_{\varepsilon_{1}}^{2}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}\right)}, \text { and }  \tag{32}\\
\lim _{n_{1} \rightarrow \infty} c_{2}= & \frac{n_{2}-\bar{n}_{2}\left(\rho, \widehat{\sigma}_{\varepsilon_{1}}^{2}, \widehat{\sigma}_{\varepsilon_{2}}^{2}\right)}{\lambda_{2} \frac{1-\rho^{2}+\widehat{\sigma}_{\varepsilon_{2}}^{2}}{1-\rho}\left(\frac{n_{2}}{\left(1+\rho+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)}-\rho \frac{\widehat{\sigma}_{\varepsilon_{1}}^{2}}{(1+\rho)\left(\left(1+\widehat{\sigma}_{\varepsilon_{1}}\right)\left(1+\widehat{\sigma}_{\varepsilon_{2}}^{2}\right)-\rho^{2}\right)}\right)} . \tag{33}
\end{align*}
$$

Next, in relation to the expressions for $d_{1}$ and $d_{2}$, we have that

$$
\lim _{n_{1} \rightarrow \infty} d_{1}=\lim _{n_{1} \rightarrow \infty} \frac{1}{\left(n_{1}-1\right) c_{1}+n_{2} c_{2}}=\lim _{n_{1} \rightarrow \infty} \frac{1}{\left(\frac{\left(n_{1}-1\right)}{n_{1}} n_{1} z+n_{2}\right) c_{2}}=\frac{1}{\left(\lim _{n_{1} \rightarrow \infty} n_{1} z+n_{2}\right) \lim _{n_{1} \rightarrow \infty} c_{2}}>0
$$

The fact that $n_{1} z$ and $c_{2}$ converge to a positive finite number (see (31) and (33)) implies that $d_{1}$ does not converge to zero (provided that $\rho \widehat{\sigma}_{\varepsilon_{1}}^{2}>0$; if $\rho \widehat{\sigma}_{\varepsilon_{1}}^{2}=0$, then it is easy to see that $\left.\lim _{n_{1} \rightarrow \infty} n_{1} z=\infty\right)$. A similar result is obtained with the limit of $d_{2}$. In particular,

$$
\lim _{n_{1} \rightarrow \infty} d_{2}=\left(\left(\lim _{n_{1} \rightarrow \infty} n_{1} z+n_{2}-1\right) \lim _{n_{1} \rightarrow \infty} c_{2}\right)^{-1}>\lim _{n_{1} \rightarrow \infty} d_{1}>0
$$

Proof of Proposition 6: Suppose that $n_{1}$ and $n_{2}$ go to infinity and that $n_{1} /\left(n_{1}+n_{2}\right)$ converges to $\mu_{1}$. Taking limits in the equation that characterizes $z$ (i.e., Equation (21)) and operating, we have

$$
\begin{equation*}
z=\frac{\Xi_{1} \Delta_{1}^{-1}+\mu_{2} \mu_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right) \lambda_{1} \lambda_{2}^{-1}}{\mu_{1} \mu_{2}^{-1}\left(1-\Xi_{1} \Delta_{1}^{-1}\right)+\Xi_{2} \Delta_{2}^{-1} \lambda_{1} \lambda_{2}^{-1}} \tag{34}
\end{equation*}
$$

Moreover, taking the limit in the expressions of the equilibrium coefficients given in Lemma A2, it follows that

$$
\begin{aligned}
b_{i} & =\frac{\Psi_{i}}{\mu_{i} \mu_{j}} \frac{\mu_{i} \Xi_{j} \frac{a_{i}}{a_{j}}-\mu_{j} \Psi_{j}}{\left(\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}\right)} q+a_{i}\left(\frac{\Xi_{j} \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}}{\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}}-\bar{\theta}_{i}\right), a_{i} c_{i}, i, j=1,2, j \neq i, \\
c_{1} & =\frac{\Xi_{1} \Delta_{1}^{-1}-\mu_{1} \mu_{2}^{-1}\left(1-\Xi_{1} \Delta_{1}^{-1}\right) z}{\lambda_{1}}, \text { and } c_{2}=\frac{\Xi_{2} \Delta_{2}^{-1}-\mu_{2} \mu_{1}^{-1}\left(1-\Xi_{2} \Delta_{2}^{-1}\right) z^{-1}}{\lambda_{2}}
\end{aligned}
$$

Substituting (34) in the previous expressions and after some algebra, we get

$$
\begin{aligned}
b_{i} & =\frac{\lambda_{j} \Psi_{i}}{\mu_{j} \lambda_{i} \Xi_{j}+\mu_{i} \lambda_{j} \Psi_{i}} q+a_{i}\left(\frac{\Xi_{j} \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}}{\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}}-\bar{\theta}_{i}\right), a_{i}=\frac{\mu_{j}\left(\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}\right)}{\mu_{j} \lambda_{i} \Xi_{j}+\mu_{i} \lambda_{j} \Psi_{i}}, \text { and } \\
c_{i} & =\frac{\mu_{j}\left(\Xi_{j}-\Psi_{i}\right)}{\mu_{j} \lambda_{i} \Xi_{j}+\mu_{i} \lambda_{j} \Psi_{i}}, i, j=1,2, j \neq i .
\end{aligned}
$$

Next, we derive the equilibrium in the following continuous setup: Consider now that there is a continuum of bidders $[0,1]$. Let $q$ denote the aggregate quantity supplied in the market. Suppose that a fraction $\mu_{1}$ of these bidders are traders of type 1 and the remainder fraction, $\mu_{2}$, are bidders of type 2 .

Consider a trader of type $i$. This bidder chooses to maximize

$$
\mathbb{E}\left[\pi_{i} \mid s_{i}, p\right]=\left(\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p\right) x_{i}-\lambda_{i} x_{i}^{2} / 2
$$

The FOC is given by $\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p-\lambda_{i} x_{i}=0$, or equivalently,

$$
\begin{equation*}
X_{i}\left(s_{i}, p\right)=\left(\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p\right) / \lambda_{i} \tag{35}
\end{equation*}
$$

Positing linear strategies, the market clearing condition implies that

$$
\begin{equation*}
p=\frac{\mu_{i}\left(b_{i}+a_{i} s_{i}\right)+\mu_{j}\left(b_{j}+a_{j} s_{j}\right)-q}{\mu_{i} c_{i}+\mu_{j} c_{j}}, \tag{36}
\end{equation*}
$$

provided that $\mu_{i} c_{i}+\mu_{j} c_{j} \neq 0$. Using the expression for $p$ and assuming that $a_{i} \neq 0, i=1,2$, it follows that $\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]=\mathbb{E}\left[\theta_{i} \mid s_{i}, s_{j}\right]$. Hence, $\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]=\bar{\theta}_{i}+\Xi_{i}\left(s_{i}-\bar{\theta}_{i}\right)+\Psi_{i}\left(s_{j}-\bar{\theta}_{j}\right)$. Using (36), $s_{j}=\left(q-\mu_{i} b_{i}-\mu_{j} b_{j}-\mu_{i} a_{i} s_{i}+p\left(\mu_{i} c_{i}+\mu_{j} c_{j}\right)\right) /\left(\mu_{j} a_{j}\right)$. Therefore,

$$
\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]=\bar{\theta}_{i}+\Xi_{i}\left(s_{i}-\bar{\theta}_{i}\right)+\Psi_{i}\left(\frac{q-\mu_{i} b_{i}-\mu_{j} b_{j}-\mu_{i} a_{i} s_{i}+p\left(\mu_{i} c_{i}+\mu_{j} c_{j}\right)}{\mu_{j} a_{j}}-\bar{\theta}_{j}\right) .
$$

Substituting this expression in (35), and identifying coefficients, it follows that

$$
\begin{align*}
b_{i} & =\left(\left(1-\Xi_{i}\right) \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}+\frac{\Psi_{i}\left(q-\left(\mu_{i} b_{i}+\mu_{j} b_{j}\right)\right)}{a_{j} \mu_{j}}\right) / \lambda_{i}  \tag{37}\\
a_{i} & =\left(\Xi_{i}-\Psi_{i} \frac{\mu_{i} a_{i}}{\mu_{j} a_{j}}\right) / \lambda_{i}, \text { and }  \tag{38}\\
c_{i} & =\left(1-\frac{\Psi_{i}}{\mu_{j} a_{j}}\left(\mu_{i} c_{i}+c_{j} \mu_{j}\right)\right) / \lambda_{i}, i, j=1,2, j \neq i . \tag{39}
\end{align*}
$$

Note that (38) implies that $a_{i} / a_{j}=\lambda_{j}\left(\Xi_{i}-\Psi_{i} \mu_{i} a_{i} /\left(\mu_{j} a_{j}\right)\right) /\left(\lambda_{i}\left(\Xi_{j}-\Psi_{j} \mu_{j} a_{j} /\left(\mu_{i} a_{i}\right)\right)\right)$. Hence, $a_{i} / a_{j}=\mu_{j}\left(\Psi_{j} \lambda_{i} \mu_{j}+\Xi_{i} \lambda_{j} \mu_{i}\right) /\left(\mu_{i}\left(\Psi_{i} \lambda_{j} \mu_{i}+\lambda_{i} \Xi_{j} \mu_{j}\right)\right)$. Then, plugging the previous expression into (38), we get

$$
\begin{equation*}
a_{i}=\mu_{j}\left(\Xi_{i} \Xi_{j}-\Psi_{i} \Psi_{j}\right) /\left(\mu_{i} \Psi_{i} \lambda_{j}+\mu_{j} \lambda_{i} \Xi_{j}\right) \tag{40}
\end{equation*}
$$

Using (37), and after some algebra, we get

$$
\mu_{i} b_{i}+\mu_{j} b_{j}=\frac{\frac{\mu_{i}}{\lambda_{i}}\left(\bar{\theta}_{i}\left(1-\Xi_{i}\right)-\Psi_{i} \bar{\theta}_{j}+q \frac{\Psi_{i}}{\mu_{j} a_{j}}\right)+\frac{\mu_{j}}{\lambda_{j}}\left(\bar{\theta}_{j}\left(1-\Xi_{j}\right)-\Psi_{j} \bar{\theta}_{i}+q \frac{\Psi_{j}}{\mu_{i} a_{i}}\right)}{\frac{\Psi_{i}}{\lambda_{i}} \frac{\mu_{i}}{\mu_{j} a_{j}}+\frac{\Psi_{j}}{\lambda_{j} \mu_{i}} \frac{\mu_{j}}{a_{i}}+1} .
$$

Substituting (40) and the last expression in (37),

$$
b_{i}=\frac{\lambda_{j} \Psi_{i}}{\mu_{i} \lambda_{j} \Psi_{i}+\mu_{j} \lambda_{i} \Xi_{j}} q+\frac{\mu_{j}\left(\Xi_{j} \Xi_{i}-\Psi_{j} \Psi_{i}\right)}{\mu_{i} \lambda_{j} \Psi_{i}+\mu_{j} \lambda_{i} \Xi_{j}}\left(\frac{\Xi_{j} \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}}{\Xi_{j} \Xi_{i}-\Psi_{j} \Psi_{i}}-\bar{\theta}_{i}\right) .
$$

Moreover, from (39), and after some algebra, we get $\mu_{i} c_{i}+c_{j} \mu_{j}=\left(\frac{\mu_{i}}{\lambda_{i}}+\frac{\mu_{j}}{\lambda_{j}}\right) /\left(\frac{\mu_{i}}{\mu_{j} a_{j}} \frac{\Psi_{i}}{\lambda_{i}}+\frac{\mu_{j}}{\mu_{i} a_{i}} \frac{\Psi_{j}}{\lambda_{j}}+1\right)$. Using (40) and the last expression in (39), it follows that $c_{i}=\mu_{j}\left(\Xi_{j}-\Psi_{i}\right) /\left(\mu_{j} \lambda_{i} \Xi_{j}+\mu_{i} \lambda_{j} \Psi_{i}\right)$. Comparing the equilibrium coefficients of the limiting case with the ones of the continuous setup, we conclude that the equilibrium coefficients of the limiting case converge to the equilibrium coefficients of the continuous setup. Finally, taking into account the expressions for $\Xi_{i}, \Xi_{j}, \Psi_{i}$, and $\Psi_{j}$, we obtain the expressions stated in Proposition 6.

Lemma A3. The equilibrium quantities solves the following distorted benefit maximization program:

$$
\begin{gathered}
\max _{x_{1}, x_{2}} \mathbb{E}\left[\left.n_{1}\left(\theta_{1} x_{1}-\left(d_{1}+\lambda_{1}\right) \frac{x_{1}^{2}}{2}\right)+n_{2}\left(\theta_{2} x_{2}-\left(d_{2}+\lambda_{2}\right) \frac{x_{2}^{2}}{2}\right) \right\rvert\, t\right] \\
\text { s.t. } n_{1} x_{1}+n_{2} x_{2}=Q
\end{gathered}
$$

taking as given the equilibrium parameters $d_{1}$ and $d_{2}$.
Proof: The Lagrangian function of the maximization program is given by

$$
\mathcal{L}\left(x_{1}, x_{2}, \mu\right)=n_{1}\left(t_{1} x_{1}-\left(d_{1}+\lambda_{1}\right) \frac{x_{1}^{2}}{2}\right)+n_{2}\left(t_{2} x_{2}-\left(d_{2}+\lambda_{2}\right) \frac{x_{2}^{2}}{2}\right)-\mu\left(n_{1} x_{1}+n_{2} x_{2}-Q\right),
$$

where $\mu$ denotes the Lagrange multiplier. Differentiating, we obtain the following FOCs:

$$
\begin{align*}
n_{1}\left(t_{1}-\left(d_{1}+\lambda_{1}\right) x_{1}\right)-\mu n_{1} & =0,  \tag{41}\\
n_{2}\left(t_{2}-\left(d_{2}+\lambda_{2}\right) x_{2}\right)-\mu n_{2} & =0, \text { and }  \tag{42}\\
n_{1} x_{1}+n_{2} x_{2} & =Q . \tag{43}
\end{align*}
$$

From (41) and (42), it follows that $x_{i}=\left(t_{i}-\mu\right) /\left(d_{i}+\lambda_{i}\right), i=1,2$. Substituting these expressions in (43) and operating, we have $\mu=\left(n_{1} \frac{t_{1}}{d_{1}+\lambda_{1}}+n_{2} \frac{t_{2}}{d_{2}+\lambda_{2}}-Q\right)\left(\frac{n_{1}}{d_{1}+\lambda_{1}}+\frac{n_{2}}{d_{2}+\lambda_{2}}\right)^{-1}$. Hence,

$$
x_{i}=\frac{n_{j}\left(t_{i}-t_{j}\right)}{n_{i}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(d_{i}+\lambda_{i}\right)}+\frac{d_{j}+\lambda_{j}}{n_{i}\left(d_{j}+\lambda_{j}\right)+n_{j}\left(d_{i}+\lambda_{i}\right)} Q, i=1,2,
$$

i.e., the equilibrium quantities. In addition, since the objective function is concave and the constraint is a linear equation, we conclude that the critical point is a global maximum. Hence, we conclude that the equilibrium quantities are the solutions of the optimization problem stated in Lemma A3.

Proof of Proposition 7: Recall that in the competitive setup, the FOC of the two optimization problems are given by $\mathbb{E}\left[\theta_{i} \mid s_{i}, p\right]-p-\lambda_{i} x_{i}=0, i=1,2$. Doing similar computations as in the proof of Lemma A1, we derive the following system of equations: ${ }^{23}$

$$
\begin{aligned}
b_{i} & =\frac{\left(1-\Xi_{i}\right) \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}+\Psi_{i}\left(\frac{Q-n_{i} b_{i}-n_{j} b_{j}}{n_{j} a_{j}}\right)}{\lambda_{i}}, a_{i}=\frac{\Xi_{i}-\frac{n_{i}}{n_{j}} \frac{a_{i}}{a_{j}} \Psi_{i}}{\lambda_{i}}, \text { and } \\
c_{i} & =\frac{1-\Psi_{i}\left(\frac{n_{i} c_{i}+n_{j} c_{j}}{n_{j} a_{j}}\right)}{\lambda_{i}}, i, j=1,2, j \neq i
\end{aligned}
$$

Taking into account that $Q=\left(n_{i}+n_{j}\right) q$ and $\mu_{i}=n_{i} /\left(n_{i}+n_{j}\right)$, we have that the previous system is identical to the system of equations given in (37)-(39). As a consequence, the equilibrium coefficients given in the statement of Proposition 6 coincide with the equilibrium coefficients in the competitive setup.

Proof of Proposition 8: Performing similar computations as in the proof of Lemma A1, we obtain that the equilibrium coefficients with subsidies $\kappa_{i}=d_{i}\left(c_{1}^{o}, c_{2}^{o}\right)$ satisfy

$$
\begin{aligned}
b_{i} & =\frac{\left(1-\Xi_{i}\right) \bar{\theta}_{i}-\Psi_{i} \bar{\theta}_{j}-\frac{\Psi_{i}\left(n_{i} b_{i}+n_{j} b_{j}-Q\right)}{n_{j} a_{j}}}{d_{i}+\lambda_{i}-d_{i}\left(c_{1}^{o}, c_{2}^{o}\right)}, a_{i}=\frac{\Xi_{i}-\frac{n_{i} a_{i}}{n_{j} a_{j}} \Psi_{i}}{d_{i}+\lambda_{i}-d_{i}\left(c_{1}^{o}, c_{2}^{o}\right)}>0, \text { and } \\
c_{i} & =\frac{1-\frac{\Psi_{i}\left(n_{i} c_{i}+n_{j} c_{j}\right)}{n_{j} a_{j}}}{d_{i}+\lambda_{i}-d_{i}\left(c_{1}^{o}, c_{2}^{o}\right)}, i, j=1,2, j \neq i .
\end{aligned}
$$

[^11]Comparing this system of equations and the one derived in the proof of Proposition 7, we obtain that the equilibrium coefficients of the price-taking equilibrium solves this system. Therefore, we conclude that the quadratic subsidies $\kappa_{i} x_{i}^{2} / 2$, with $\kappa_{i}=d_{i}\left(c_{1}^{o}, c_{2}^{o}\right), i=1,2$, induce an efficient allocation.

Lemma A4. The expected deadweight loss at an anonymous allocation $\left(x_{1}(t), x_{2}(t)\right)$ satisfies

$$
\begin{equation*}
\mathbb{E}[D W L]=\frac{1}{2} \lambda_{1} n_{1} \mathbb{E}\left[\left(x_{1}(t)-x_{1}^{o}(t)\right)^{2}\right]+\frac{1}{2} \lambda_{2} n_{2} \mathbb{E}\left[\left(x_{2}(t)-x_{2}^{o}(t)\right)^{2}\right] . \tag{44}
\end{equation*}
$$

Proof: Notice that $E T S=\mathbb{E}[\mathbb{E}[T S \mid t]]$, where

$$
\begin{aligned}
\mathbb{E}[T S \mid t]= & \mathbb{E}\left[n_{1}\left(\theta_{1} x_{1}(t)-\lambda_{1}\left(x_{1}(t)\right)^{2} / 2\right)+n_{2}\left(\theta_{2} x_{2}(t)-\lambda_{2}\left(x_{2}(t)\right)^{2} / 2\right) \mid t\right]= \\
& n_{1}\left(t_{1} x_{1}(t)-\lambda_{1}\left(x_{1}(t)\right)^{2} / 2\right)+n_{2}\left(t_{2} x_{2}(t)-\lambda_{2}\left(x_{2}(t)\right)^{2} / 2\right)
\end{aligned}
$$

A Taylor series expansion of $\mathbb{E}[T S \mid t]$ around the price-taking equilibrium $\left(x_{1}^{o}(t), x_{2}^{o}(t)\right)$, stopping at the second term due to the fact that $\mathbb{E}[T S \mid t]$ is quadratic, yields

$$
\begin{aligned}
\mathbb{E}[T S \mid t](x(t))= & \mathbb{E}[T S \mid t]\left(x^{o}(t)\right)+\nabla \mathbb{E}[T S \mid t]\left(x^{o}(t)\right)\left(x(t)-x^{o}(t)\right)+ \\
& +\frac{1}{2}\left(x(t)-x^{o}(t)\right)^{\prime} D^{2} \mathbb{E}[T S \mid t]\left(x^{o}(t)\right)\left(x(t)-x^{o}(t)\right),
\end{aligned}
$$

where $\nabla \mathbb{E}[T S \mid t]\left(x^{o}(t)\right)$ and $D^{2} \mathbb{E}[T S \mid t]\left(x^{o}(t)\right)$ are, respectively, the gradient and the Hessian matrix of $\mathbb{E}[T S \mid t]$ evaluated at $x^{o}(t)$. Notice that we know

$$
\nabla \mathbb{E}[T S \mid t]\left(x^{o}(t)\right)=\left(n_{1}\left(t_{1}-\lambda_{1} x_{1}^{o}(t)\right), n_{2}\left(t_{2}-\lambda_{2} x_{2}^{o}(t)\right)\right) .
$$

Using the expressions of $x_{1}^{o}(t)$ and $x_{2}^{o}(t)$,

$$
\begin{aligned}
& \nabla \mathbb{E}[T S \mid t]\left(x^{o}(t)\right)\left(x(t)-x^{o}(t)\right)=\left(n_{1}\left(t_{1}-\lambda_{1} x_{1}^{o}(t)\right), n_{2}\left(t_{2}-\lambda_{2} x_{2}^{o}(t)\right)\right)\binom{x_{1}(t)-x_{1}^{o}(t)}{x_{2}(t)-x_{2}^{o}(t)}= \\
& \left(\frac{n_{1} \frac{t_{1}}{\lambda_{1}}+n_{2} \frac{t_{2}}{\lambda_{2}}}{\frac{n_{1}}{\lambda_{1}}+\frac{n_{2}}{\lambda_{2}}}-\frac{Q}{\frac{n_{1}}{\lambda_{1}}+\frac{n_{2}}{\lambda_{2}}}\right)\left(n_{1}\left(x_{1}-x_{1}^{o}\right)+n_{2}\left(x_{2}-x_{2}^{o}\right)\right)=\left(\frac{n_{1} \frac{t_{1}}{\lambda_{1}}+n_{2} \frac{t_{2}}{\lambda_{2}}}{\frac{n_{1}}{\lambda_{1}}+\frac{n_{2}}{\lambda_{2}}}-\frac{Q}{\frac{n_{1}}{\lambda_{1}}+\frac{n_{2}}{\lambda_{2}}}\right)(Q-Q)=0 . \\
& \text { In addition, } D^{2} \mathbb{E}[T S \mid t]\left(x^{o}(t)\right)=\left(\begin{array}{cc}
-\lambda_{1} n_{1} & 0 \\
0 & -\lambda_{2} n_{2}
\end{array}\right) \text {. Hence, } \\
& \mathbb{E}[T S \mid t](x(t))-\mathbb{E}[T S \mid t]\left(x^{o}(t)\right)=-\frac{1}{2} \lambda_{1} n_{1}\left(x_{1}(t)-x_{1}^{o}(t)\right)^{2}-\frac{1}{2} \lambda_{2} n_{2}\left(x_{2}(t)-x_{2}^{o}(t)\right)^{2}
\end{aligned}
$$

and, therefore, (44) is satisfied.
Proof of Proposition 9: i) Suppose that $Q=0$. Then, $\mathbb{E}[D W L]$ is given by

$$
\mathbb{E}[D W L]=\frac{n_{2} n_{1}\left(n_{2} d_{1}+n_{1} d_{2}\right)^{2}}{2\left(n_{2} \lambda_{1}+n_{1} \lambda_{2}\right)\left(n_{2}\left(d_{1}+\lambda_{1}\right)+n_{1}\left(d_{2}+\lambda_{2}\right)\right)^{2}} \mathbb{E}\left(t_{1}-t_{2}\right)^{2}
$$

Hence, $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} \mathbb{E}[D W L]=\frac{\partial}{\partial d_{1}} \mathbb{E}[D W L] \frac{d}{d \sigma_{\varepsilon_{1}}^{2}} d_{1}+\frac{\partial}{\partial d_{2}} \mathbb{E}[D W L] \frac{d}{d \sigma_{\varepsilon_{1}}^{2}} d_{2}+\frac{\partial}{\partial \sigma_{\varepsilon_{1}}^{2}} \mathbb{E}[D W L]$. It is easy to see that in this case $\frac{\partial}{\partial d_{i}} \mathbb{E}[D W L]>0, i=1,2$, and $\frac{\partial}{\partial \sigma_{\varepsilon_{1}}^{2}} \mathbb{E}[D W L]<0$. Combining these results with Proposition 3, we have that the first two terms of $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} \mathbb{E}[D W L]$ are positive, while the last one is negative.

We know that $d_{1}$ and $d_{2}$ are independent of $\sigma_{\varepsilon_{1}}^{2}$ when $\rho=0$. By continuity, we know that for very low values of $\rho$ is $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} d_{1}$ and $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} d_{2}$ are positive, but very low. Hence, we conclude that the last term in $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} \mathbb{E}[D W L]$ dominates and, hence, in this case we have that $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} \mathbb{E}[D W L]<0$ although $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} d_{i}>0, i=1,2$. By contrast, if we consider the case in which $\rho$ is not low and $\left(\bar{\theta}_{1}-\bar{\theta}_{2}\right)^{2}$ is high enough, then the first two terms in $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} \mathbb{E}[D W L]$ dominate, which implies that $\frac{d}{d \sigma_{\varepsilon_{1}}^{2}} \mathbb{E}[D W L]>0, i=1,2$.
ii) Omitted since it is trivial.
iii) When groups are symmetric $\left(n_{2}=n_{1}=n, \lambda_{1}=\lambda_{2}=\lambda\right.$, and $\left.\sigma_{\varepsilon_{1}}^{2}=\sigma_{\varepsilon_{2}}^{2}=\sigma_{\varepsilon}^{2}\right)$ $d_{1}=d_{2}=d$ and $\lambda_{2} d_{1}-\lambda_{1} d_{2}=0$. Therefore, the expected deadweight loss consists of only one term, the first one, that is independent of $Q$. Moreover, after some computations, we get that $\mathbb{E}[D W L \mid t]=\frac{n d^{2}}{4(d+\lambda)^{2} \lambda}\left(t_{2}-t_{1}\right)^{2}$. Using this expression and the fact that $\frac{\partial}{\partial d} \mathbb{E}[D W L \mid t]>0$ allows us to conclude that an increase in an information parameter ( $\rho$ or $\widehat{\sigma}_{\varepsilon}^{2}$ ) raises both $d$ and $\mathbb{E}[D W L \mid t]$.

However, with asymmetric groups the previous results may not hold. In this case, suppose that $Q$ is large enough. Then, for $i=1,2, j \neq i$,

$$
\begin{aligned}
\operatorname{sign}\left(\frac{\partial}{\partial d_{i}} E[D W L \mid t]\right) & =\operatorname{sign}\left(\frac{\partial}{\partial d_{i}}\left(\frac{1}{2} n_{1} n_{2} \frac{\left(\lambda_{1} d_{2}-\lambda_{2} d_{1}\right)^{2}}{\left(\lambda_{1} n_{2}+\lambda_{2} n_{1}\right)\left(n_{1}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(d_{1}+\lambda_{1}\right)\right)^{2}} Q^{2}\right)\right) \\
& =\operatorname{sign}\left(-Q^{2} n_{1} n_{2}\left(\lambda_{j}+d_{j}\right) \frac{\lambda_{i} d_{j}-\lambda_{j} d_{i}}{\left(n_{1}\left(d_{2}+\lambda_{2}\right)+n_{2}\left(d_{1}+\lambda_{1}\right)\right)^{3}}\right)
\end{aligned}
$$

When $Q$ is large enough we have that if $\frac{d_{i}}{d_{j}}>\frac{\lambda_{i}}{\lambda_{j}}$, then $\frac{\partial}{\partial d_{i}} E[D W L \mid t]>0$ and $\frac{\partial}{\partial d_{j}} E[D W L \mid t]<$ 0 . Thus, with asymmetric groups market power $\left(d_{1}, d_{2}\right)$ and the $E[D W L \mid t]$ are not always positively associated, given predicted values $t$, for changes in information parameters.

## References

Akgün, U. (2004): "Mergers with supply functions," Journal of Industrial Economics, 52 (4), 535-546.

Anderson, E. AND X. Hu (2008): "Finding supply function equilibria with asymmetric firms," Operations Research, 56 (3), 697-711.

Armantier, O., AND E. Sbaï (2006): "Estimation and comparison of treasury auction formats when bidders are asymmetric," Journal of Applied Econometrics, 21, 745-779.

Arnone, M., AND G. Iden (2003): "Primary dealers in government securities: policy issues and selected countries' experience," IMF Working Paper 03/45.

Athey, S. (2001): "Single crossing properties and the existence of pure strategy equilibria in games of incomplete information," Econometrica, 69 (4), 861-890.

Athey, S., D. Coey AND J. Levin (2013): "Set-asides and subsidies in auctions," American Economic Journal: Microeconomics, 5 (1), 1-27.
Ausubel, L. M., P. Cramton, M. Pycia, M. Rostek, AND M. Weretka (2014): "Demand reduction and inefficiency in multi-unit auctions," The Review of Economic Studies, 81 (4), 1366-1400.

Back, K., AND J.F. Zender (1993): "Auctions of divisible goods: on the rationale for the Treasury experiment," The Review of Financial Studies, 6, 733-764.

Bajari, P. (1998): "A structural econometric model of the first price sealed bid auction with asymmetric bidders," Manuscript, Stanford University.

Bergemann, D., T. Heumann, AND S. Morris (2015): "Information and market power," Cowles Foundation working paper no. 2017.
Bikhchandani, S., AND C. Huang (1993): "The Economics of Treasury Securities Markets," Journal of Economic Perspectives, 7 (3): 117-134.

Bindseil, U., K. Nyborg, AND I. Strebulaev (2009): "Bidding and performance in repo auctions: Evidence from ECB open market operations," Journal of Money, Credit and Banking, 41 (7), 1391-1421.

Bjonnes, G.H. (2001): "Winner's curse in discriminatory price auctions: Evidence from the Norwegian Treasury bill auctions," SIFR Research Report Series 3, Institute for Financial Research.

Borenstein, S., J. Bushnell, AND F. Wolak (2002): "Measuring market inefficiencies in California's restructured wholesale electricity market," American Economic Review, 92 (5), 1376-1405.
Boyarchenko, N., D.O. Lucca, AND L. Veldkamp (2015): "Intermediaries as information aggregators: an application to US Treasury Auctions," Federal Reserve Bank of New York Staff Report no. 726.
Brandts, J., Reynolds, S. AND A. Schram (2014): "Pivotal suppliers and market power in experimental supply function competition," Economic Journal, 124 (579), 887-916.

Brenner, M., D. Galai, AND O. Sade (2009): "Sovereign debt auctions: Uniform or discriminatory?," Journal of Monetary Economics, 56 (2), 267-274.

Bustos-Salvagno, J. (2015): "Bidding behavior in the Chilean electricity market," Energy Economics, 51, 288-299.

Cantillon, E. (2008): "The effect of bidders' asymmetries on expected revenue in auctions," Games and Economic Behavior, 62 (1), 1-25.

Cassola, N., A. Hortaçsu, AND J. Kastl (2013): "The 2007 subprime market crisis through the lens of European Central Bank auctions for short-term funds," Econometrica, 81 (4), 1309-1345. Ciarreta, A., AND M.P. Espinosa (2010): "Supply function competition in the Spanish wholesale electricity market," Energy Journal, 31 (4), 137-157.

Cramton, P., AND S. Stoft (2006): "Uniform price auctions in electricity markets," U. Maryland Working paper.

Cramton, P., AND S. Stoft (2007): "Why we need to stick with uniform price auctions in electricity markets," Electricity Journal, 20 (1), 26-37.

Crawford, G., J. Crespo, AND H. Tauchen (2007): "Bidding asymmetries in multi-unit auctions: Implications of bid function equilibria in the British spot market for electricity," International Journal of Industrial Organization, 25 (6), 1233-1268.

Du, S. AND H. Zhu (2015): "Welfare and optimal trading frequency in dynamic double auctions," Working paper.

Du, S. AND H. Zhu (2016): "Bilateral trading in divisible double auctions," Working paper.
Engelbrecht-Wiggans, R., J.A. List, AND D. Lucking-Reiley (2006): "Demand reduction in multi-unit auctions with varying numbers of bidders: Theory and field experiments," International Economic Review, 47, 203-231.

European Commission (2007): "DG Competition report on energy sector inquiry," 10 January 2007, SEC (2006) 1724, Brussels, Commission of the European Communities.

Green, R., AND D. Newbery (1992): "Competition in the British Electricity Spot Market," Journal of Political Economy, 100, 929-953.

Grossman, S., AND J. Stiglitz (1980): "On the impossibility of informationally efficient markets," American Economic Review, 70, 393-408.

Holmberg, P., D. Newbery, AND D. Ralph (2013): "Supply function equilibria: Step functions and continuous representations," Journal of Economic Theory, 148(4), 1509-1551.

Holmberg, P., AND F. Wolak (2015): "Electricity markets: Designing auctions where suppliers have uncertain costs," IFN working paper no. 1099.

Hortaçsu, A., AND J. Kastl (2012): "Valuing dealers' informational advantage: a study of Canadian Treasury auctions," Econometrica, 80 (6), 2511-2542.

Hortaçsu, A., AND D. McAdams (2010): "Mechanism choice and strategic bidding in divisible good auctions: An empirical analysis of the Turkish Treasury auction market," Journal of Political Economy, 118 (5), 833-865.
Hortaçsu, A., AND S. Puller (2008): "Understanding strategic bidding in multi-unit auctions: A case study of the Texas electricity spot market," RAND Journal of Economics, 39 (1), 86-114. Joskow, P., AND E. Kahn (2002): "A quantitative analysis of pricing behavior in California's wholesale electricity market during summer 2000," The Energy Journal, 23 (4), 1-35.
Kagel, J.H., AND D. Levin (2001): "Behavior in multi-unit demand auctions: Experiments with uniform price and dynamic Vickrey auctions," Econometrica, 69, 413-454.
Kastl, J. (2011): "Discrete bids and empirical inference in divisible good auctions," The Review of Economic Studies, 78 (3), 974-1014.
Klemperer, P., AND M. Meyer (1989): "Supply Function Equilibria in Oligopoly Under Uncertainty," Econometrica, 57, 1243-1277.

Kyle, A. (1989): "Informed speculation with imperfect competition," The Review of Economic Studies, 56, 317-356.

Lebrun, B. (1996): "Existence of an equilibrium in first price auctions," Economic Theory, 7 (3), 421-443.

Lebrun, B. (1999): "First price auctions in the asymmetric N bidder case," International Economic Review, 40 (1), 125-142.

Loertscher, S. and L. Marx (2016): "Auctions with Bid Credits and Resale", WP.
Lopomo, G., L. Marx, D. McAdams, AND B. Murray (2011): "Carbon allowance auction design: an assessment of options for the United States," Review of Environmental Economics and Policy, 5 (1) 25-43.
Malvey, P.F., AND C.F. Archibald (1998): "Uniform-price auctions: Update of the Treasury experience," US Treasury.
Maskin, E., AND J.G. Riley (2000a): "Asymmetric auctions," The Review of Economic Studies, 67 (3), 413-438.
Maskin, E., AND J.G. Riley (2000b): "Equilibrium in sealed high bid auctions," The Review of Economic Studies, 67 (3), 439-454.

Maskin, E., AND J.G. Riley (2003): "Uniqueness in sealed high bid auctions," Games and Economic Behavior 45 (2), 395-409.
McAdams, D. (2003): "Isotone equilibrium in games of incomplete information," Econometrica 71 (4), 1191-1214.

McAdams, D. (2006): "Monotone equilibrium in multi-unit auctions," Review of Economic Studies 73, 1039-1056.

Pesendorfer, M. (2000): "A study of collusion in first-price auctions," The Review of Economic Studies, 67 (3), 381-411.

Pai, M. and R. Vohra (2012): "Auction Design with Fairness Concerns: Subsidies vs. SetAsides," Discussion Paper, Center for Mathematical Studies in Economics and Management Science, Northwestern University.

Porter, R.H., AND J.D. Zona (1999): "Ohio school milk markets: An analysis of bidding," RAND Journal of Economics, 30 (2), 263-288.

Pycia, M., AND K. Woodward (2016): "Pay-as-bid. Selling divisible goods," Working paper.
Reny, P. (2011): "On the existence of monotone pure-strategy equilibria in Bayesian games," Econometrica 79 (2), 499-553.

Reny, P., AND S. Zamir (2004): "On the existence of pure strategy monotone equilibria in asymmetric first-price auctions," Econometrica, 72 (4), 1105-1125.

Rostek, M., AND M. Weretka (2012): "Price inference in small markets," Econometrica 80 (2), 687-711.

Sadzik, T., AND P. Andreyanov (2016): "Robust Mechanism Design of Exchange," Working paper.

Umlauf, S.R. (1993): "An empirical study of the Mexican Treasury bill auction," Journal of Financial Economics, 33, 313-340.

Vives, X. (2010): "Asset auctions, information, and liquidity," Journal of the European Economic Association, 8 (2-3), 467-477.

Vives, X. (2011): "Strategic supply function competition with private information," Econometrica, 79 (6), 1919-1966.

Vives, X. (2014): "On the possibility of informationally efficient markets," Journal of the European Economic Association, 12 (5), 1200-1239.

Wang, J.J.D., AND J.F. Zender (2002): "Auctioning divisible goods," Economic Theory, 19 (4), 673-705.

Wilson, R. (1979): "Auctions of shares," Quarterly Journal of Economics, 93, 675-698.
Wolfram, C. (1998): "Strategic Bidding in a Multiunit Auction: An Empirical Analysis of Bids to Supply Electricity in England and Wales," RAND Journal of Economics, 29, 703-725.


[^0]:    Terms of use:
    Documents in EconStor may be saved and copied for your personal and scholarly purposes.

    You are not to copy documents for public or commercial purposes, to exhibit the documents publicly, to make them publicly available on the internet, or to distribute or otherwise use the documents in public.

    If the documents have been made available under an Open Content Licence (especially Creative Commons Licences), you may exercise further usage rights as specified in the indicated licence.

[^1]:    ${ }^{1}$ See Lopomo et al. (2011) for examples of such auctions.
    ${ }^{2}$ The relatively small number of primary dealers makes the US Treasury market imperfectly competitive (Bikhchandani and Huang 1993). Uniform-price auctions are often used in Treasury, liquidity and electricity auctions, for example. See Brenner et al. (2009) for Treasury auctions, with the United States a leading example since November 1998. Experimental work has found substantial demand reduction in uniform-price auctions (see e.g. Kagel and Levin 2001; Engelbrecht-Wiggans et al. 2006).

[^2]:    ${ }^{3}$ Bidder asymmetry has also been found in procurement markets, including school milk (Porter and Zona 1999; Pesendorfer 2000) and public works (Bajari 1998).
    ${ }^{4}$ European Commission (2007) has asserted that "at the wholesale level, gas and electricity markets remain national in scope, and generally maintain the high level of concentration of the pre-liberalization period. This gives scope for exercising market power" (Inquiry pursuant to Article 17 of Regulation (EC) No 1/2003 into the European gas and electricity sectors (Final Report), Brussels, 10.1.2007).
    ${ }^{5}$ One reason for differences in private information among bidders may be the presence of both dealers and direct bidders in auctions (such as in US Treasury auctions). Dealers aggregate the information of clients and bid with a higher precision of information (for evidence from Canadian Treasury auctions, see Hortaçsu and

[^3]:    ${ }^{8}$ Wilson (1979) compares a uniform-price auction for a divisible good with an auction in which the good is treated as an indivisible good; he finds that the price can be significantly lower if bidders are allowed to submit bid schedules rather than a single bid price. That work is extended by Back and Zender (1993), who compare a uniform-price auction with a discriminatory auction. These authors demonstrate the existence of equilibria in which the seller's revenue in a uniform-price auction can be much lower than the revenue obtained in a discriminatory auction. According to Wang and Zender (2002), if supply is uncertain and bidders are risk averse, then there may exist equilibria of a uniform-price auction that yield higher expected revenue than that from a discriminatory auction.
    ${ }^{9}$ Ausubel et al. (2014) find that, in symmetric auctions with decreasing linear marginal utility, the seller's revenue is greater in a discriminatory auction than in a uniform-price auction. Pycia and Woodward (2016) demonstrate that a discriminatory pay-as-bid auction is revenue-equivalent to the uniform-price auction provided that supply and reserve prices are set optimally.
    ${ }^{10}$ This assumption states that the sum of correlations in each column of this matrix (or, equivalently, in each row) is the same and that the variances of all traders' values are also the same. Unlike our model, Rostek and Weretka's (2012) model maintains the symmetry assumption as regards transaction costs and the precision

[^4]:    ${ }^{12}$ The value of $\rho$ will depend of the type of security. In this sense, Bindseil et al. (2009) argue that the common value component is less important in a central bank repo auction than in a T-bill auction.

[^5]:    ${ }^{13}$ We assume that $\left(n_{i}-1\right) c_{i}+n_{j} c_{j} \neq 0$.

[^6]:    ${ }^{14}$ In the full (shared) information setup, traders can access $\left(s_{1}, s_{2}\right)$. In this framework the price does not provide any useful information.
    ${ }^{15} \mathrm{Du}$ and Zhu (2016) consider ex post nonlinear equilibria in a bilateral divisible double auction and show that with more than three symmetric traders there are no nonlinear equilibria in the class of smooth demands downward sloping in price and upward sloping in signals.

[^7]:    ${ }^{16}$ This result (in the supply competition model) may help explain why, in the Texas balancing market, small firms use steeper supply functions than predicted by theory (Hortaçsu and Puller 2008). Indeed, smaller firms may receive lower-quality signals owing to economies of scale in information gathering.

[^8]:    ${ }^{17}$ A high price conveys the good news that the private signal received by other group's traders is high. When valuations are positively correlated, a bidder infers from the high private signal of the other group that her own valuation is high.
    ${ }^{18}$ This claim follows because if $n_{1}=n_{2}$ and $\widehat{\sigma}_{\varepsilon_{1}}^{2}=\widehat{\sigma}_{\varepsilon_{2}}^{2}$, then the inequality given in (4) does not hold.

[^9]:    ${ }^{19}$ See Lemma A3 in the Appendix.

[^10]:    ${ }^{20}$ The expected optimal subsidy for group $i$ is $\kappa_{i} \mathbb{E}\left[\left(x_{i}^{o}(t)\right)^{2}\right] / 2$, where $x_{i}^{o}(t)=$ $\left(n_{j}\left(t_{i}-t_{j}\right)+\lambda_{j} Q\right) /\left(n_{i} \lambda_{j}+n_{j} \lambda_{i}\right)$.
    ${ }^{21}$ Athey et al. (2013) find with regard to US Forest Service timber auctions that restricting entry increases small business participation but substantially reduces efficiency and revenue. In contrast, subsidizing small bidders directly increases revenue and the profits of small bidders without much cost in efficiency. See also Loertscher and Marx (2016) and Pai and Vohra (2012).

[^11]:    ${ }^{23}$ To ease the notation the superscript $o$ is omitted in this proof.

