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The Lasting Effect of Sex Ratio Imbalance on Marriage and Family: Evidence from World War II in Russia

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ABSTRACT

The Lasting Effect of Sex Ratio Imbalance on Marriage and Family: Evidence from World War II in Russia*

How does a shock to sex ratios affect marriage markets and fertility? I use the drastic change in sex ratios caused by World War II to identify the effects of unbalanced sex ratios on Russian women. Using unique data from the Soviet archives, the results indicate that male scarcity led to lower rates of marriage and fertility, higher nonmarital births and reduced bargaining power within marriage for women most affected by war deaths. The impact of sex ratio imbalance on marriage and family persisted for years after the war's end, and was likely magnified by policies that promoted nonmarital births and discouraged divorce.

JEL Classification: J12, J16, N34, P23

Keywords: sex ratios, marriage, family, fertility, divorce law

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In many ways the Soviet Union was built on the backs of Russian women. Pulled into the labor force by the intense industrialization drive of the 1930s and indispensable for fueling the Soviet war machine of the 1940s, the experience of Russian women of the twentieth century was profoundly different from that of women in most industrialized countries. These differences are evident in the high female labor force participation rates, low fertility rates, and strikingly high abortion rates recorded in the Soviet Union over its history. Yet due to limited data, much remains unknown about the changing lives of Russian women over the twentieth century and how women's lives were affected by the many upheavals that marked the country during this turbulent period.

This paper examines the effect of one cataclysmic event, the massive loss of life in World War II, on the subsequent marital and fertility careers of Russian women. For women in the cohorts most affected by the war, these losses resulted in extremely unbalanced sex ratios in the population: for women age 20-29, for example, the ratio of men to women in the population fell from .91 in 1941 to .65 in 1946. Using unpublished Census and vital statistics data collected from the Soviet archives in Moscow, this paper uses this large, exogenous change in the sex ratio to identify the effects of highly unbalanced sex ratios on the marital and fertility outcomes of Russian women in the postwar period. Despite the elapsed years between the end of World War II and the first postwar census of 1959, the results indicate that the relative scarcity of men continued to profoundly affect women's lives: women were less likely to marry, more likely to give birth out of wedlock, and more likely to be divorced in the birth cohorts and regions facing the greatest shortages of men. The unbalanced sex ratios also affected intrahousehold bargaining power among married couples by influencing marital fertility and by increasing the share of marriages with a large age gap between bride and groom. The unique institutional context of the

postwar Soviet Union, featuring significant increases in the relative cost of marriage for men, likely exacerbated the effects of male shortage by further reducing men's propensity to marry.

The results of this paper underscore the conclusions of a growing body of research demonstrating that conditions in the marriage market have profound and long-lasting effects on individuals and society as a whole. Recent research in this literature has explored the impact of high male incarceration rates on women (Charles & Luoh, 2010; Mechoulam, 2011), the influx of mainland Chinese to Taiwan in 1949 (Francis, 2011), and the effect of unbalanced sex ratios in the U.S. immigrant population (Angrist, 2002; LaFortune, 2013). The results of these papers generally support the predicted effects of sex ratio imbalance on marriage markets and intrahousehold bargaining power as first developed by Becker (1973, 1974, 1981):¹ in conditions of male scarcity, women are less likely to marry, more likely to have children out of wedlock, more likely to work, and more likely to invest in their human capital. The cultural attitudes fostered by significant sex ratio imbalance, such as attitudes towards women's work outside the home, appear to persist over long periods of time (Grosjean & Khattar 2014).

Two recent papers have explored the impact of male scarcity due to wars on marriage markets and fertility. Abramitzky et al. (2011) examine the impact of male deaths in World War I in France on the postwar marriage market. Using Census and marriage-level data for before and after the war, the authors demonstrate a significant impact of low sex ratios on marriage and fertility, in particular that men use their advantageous position in the marriage market to 'marry up' and marry women of a higher social class. These findings illustrate that an additional mechanism through which marriage markets can adjust to sex imbalances is through a change in

¹See also Guttentag & Secord (1983), Grossbard-Shechtman (1993) and Chiappori et al. (2002) for contributions in this literature. Other papers examine the effect of unbalanced sex ratios on savings rates (Wei & Zhang, 2011) and crime (Edlund et al., 2013).

assortative matching. Bethmann & Krasickova (2013) analyze how the loss of men in World War II affected out of wedlock births in Bavaria. Using county-level data and instrumenting for the sex ratio using male war deaths, the results indicate that regions with greater loss of male life experienced significantly higher shares of out of wedlock births after the war.

This paper adds to this literature by analyzing the effect of World War II in Russia, a country that has not been previously studied in this context and which suffered large-scale losses during the war. To overcome the possible endogeneity between sex ratios and marital outcomes, I use sex ratios that vary by year of birth within Russian oblasts,² which allows me to exploit the sharp discontinuities in the sex ratio by age and to control for the many potential confounding factors which differ across oblasts but are unlikely to vary by year of birth within oblasts. This paper also highlights the importance of a country's institutional context for understanding the impact of sex ratio imbalance on marriage markets and family formation: the ease or difficulty of divorce, combined with any legal obligations to support children fathered out of wedlock, may influence whether male scarcity results in increased or decreased commitment to marriage and family. This point has received little attention in the literature on sex ratios and marriage markets.

I. The impact of World War II on the Soviet population

On June 22, 1941, Hitler's *Wehrmacht* invaded the Soviet Union to initiate what would become one of the most brutal and costly wars between two countries in history. The surprise attack on the woefully unprepared Red Army led to devastating losses for the Soviet Union in the

²The term 'oblast' refers to the territorial subdivisions of the Soviet republics, called *oblasts*, *krais*, or autonomous republics in Russian, which are roughly equivalent to U.S. states. 'Region' is used interchangeably with 'oblast' in the text. This is distinct from the 'large regions' shown in Figure 2 (North, Northwest, etc.) which are similar to U.S. Census divisions.

early phase of the war: within the first six months, the Red Army had lost nearly 5 million men – the size of the Soviet Union’s entire prewar army – and had lost territory equal to the size of the United States between the East Coast and Springfield, Illinois (Glantz, 2005).

The Soviet Union mobilized all of its resources in its subsequent fight for survival and victory. The need for manpower dictated a significant loosening of the age and nationality restrictions on conscription; men under age 18 and over 55 were conscripted into the Red Army, with residents of all fifteen Soviet republics required to serve (Glantz, 2005). Including individuals serving at the beginning of the war, 34.5 million people were drafted into the armed forces during the war, of which nearly 8.7 million died in combat (Krivosheev et al., 1997).

The total losses sustained by the Soviet Union during World War II remain a topic of controversy among scholars, and an exact accounting of the deaths may never be known. The most reliable figures are those of an expert commission which estimated the total excess deaths at 26 to 27 million or roughly 13.5 percent of the prewar population (Andreev et al., 1990). The losses in the Russian republic were similar: approximately 13.6 million died, or 12.3 percent of the 1941 population. To put these losses in context, the total civilian and military casualties in Germany during World War II were 5-7 million (6-9 percent of the 1939 population), while France and the U.K. lost less than 2 percent of their prewar population in the war.

Although people of all ages fell victim to the war, whether due to death in military operations, at the hands of occupiers, or due to the widespread undernutrition and disease that accompanied the war, the wartime casualties were nevertheless heavily concentrated among young men; an estimated 20 million of the 26-27 million excess deaths were male (Ellman & Maksudov, 1994). Russian demographers calculate that the probability of surviving between 1941 and 1946 for men age 25-34 fell from .96 – the probability in the absence of the war based

on 1940 mortality rates – to .61, and estimate that women age 20-39 outnumbered men in the Soviet Union by 10.2 million in 1946 (Andreev et al., 1993). The resulting decrease in the sex ratio is illustrated in Figure 1, which plots the ratio of men to women by single year of age in 1959 in Russia. The cohorts most affected by the war were in their teens to early twenties at the beginning of the war, and extended to include those in their thirties, born in the first decades of the twentieth century. While women in the prewar Soviet Union already contended with sex ratios below 1.0 due to the 1917 Revolution, World War I and Civil War, the sex ratio fell dramatically for individuals born in the 1920s and reached a low of 0.60 for women born in 1924. The sex ratio returned to approximately 1.0 for those born in 1935 and after.

Along with its age and sex structure, the war changed the regional distribution of the population in the USSR. The western regions of the country experienced most of the fighting and sustained the greatest war losses. These losses, combined with the evacuation of millions of people eastward, led to a shift of nearly 4 percent of the population from western to eastern regions of the Soviet Union between 1939 and 1951 (Rowland, 1997). Over 1,500 factories were also relocated to east of the Volga River during the war (Barber & Harrison, 1991).

The profound impact of the war on the age and sex structure of the population across Russia's regions is illustrated in Figure 2, which shows sex ratios by region in 1959 for individuals age 25-29 and 35-39. Sex ratios for the 35-39 year-old cohort are around 0.60 in most regions, as compared with approximately 1.0 for the age 25-29 cohort that was largely unaffected by war deaths. Sex ratios for both groups are highest in the North, East Siberia and Far East, which attract disproportionately male workers to work in the natural resource sectors of the economy located in those regions. Although the occupation and physical destruction of the country occurred in the west, Figure 2 illustrates that male deaths (as proxied by the sex ratio)

occurred across all regions of the country in roughly similar proportion. This is also true of the other Soviet republics; see Online Appendix Table 1 which provides sex ratios by Soviet republic and RSFSR region for the 25-29 and 35-39 age groups.

Besides the massive loss of civilian and military lives during the war, an additional demographic cost was the decline in births during the war years. The birth rate fell from 34.6 per 1,000 population in 1940 to 26.0 in 1946; an estimated 11.5 million births did not occur because of the war (Ellman & Maksudov, 1994). Demobilization after the war took three years, further delaying the return to normal family-formation patterns until the late 1940s. This is evident in the sharp increase in age at first marriage for Russian women in the cohorts most affected by the war (see Figure 3). Despite late marriage, these women appear to have completed fertility similar to that of women in nearby cohorts. Some demographers argue that the war ultimately had little impact on the marital and fertility careers of Russian women as most women eventually married and had two children (Scherbov & Van Vianen, 2001). The analysis presented below suggests in contrast that the unbalanced sex ratios did affect fertility and family formation, as well as out of wedlock births, the spousal age gap and marital stability.

Alarmed at the devastating population losses suffered by the country and the declining birth rate, the Soviet government implemented a strongly pronatalist Family Code in 1944. This legislation imposed a tax on single people and married couples with fewer than three children, and expanded the child benefit program to provide a monthly payment for all children born out of wedlock (Heer, 1977). Far from discouraging nonmarital births, the 1944 law absolved fathers of any financial or legal responsibility for children fathered outside of marriage; unmarried mothers were prohibited from naming the father on the birth certificate or claiming financial support for their children. The 1944 Family Code also made the procedure for divorce so expensive and

complicated that it has been described as effectively a “prohibition on divorce” (Avdeev & Monnier, 2000). The high cost of divorce combined with nearly costless nonmarital sexual relations significantly increased the cost of registered marriage relative to bachelorhood for men. While Soviet leaders promoted the 1944 Family Code as a means to strengthen the family, the unique institutional context created by these laws likely exacerbated the effect of male shortage on male behavior and women’s outcomes, reducing registered marriage and increasing marital instability. These effects are investigated in the statistical analysis below.

II. Data sources

The primary empirical strategy exploits the differences in the sex ratio by year of birth and region to identify the effect of unbalanced sex ratios on marriage and fertility in Russia. The data used in the analysis are unpublished data from the 1959 Soviet Census and the 1959 Soviet vital statistics registry, collected and digitized by the author from records in the Soviet archives. The Census data comprise the distribution of the population by sex, marital status, and urban/rural area by single year of age for each of the 75 regions of the Russian Soviet Federal Socialist Republic (RSFSR).³ The vital statistics data include the number of births by marital status in each region, the number of marriages by age of bride and groom, the number of divorces, and the number of deaths due to abortion in 1959; most of the vital statistics data are given by 5-year age group as detailed in Table 1.⁴ These data are combined with the Census

³‘Marital status’ indicates whether an individual is currently married; no information is available on divorce, previous marriages, or widowhood. The terms ‘RSFSR’, ‘Russian republic,’ and ‘Russia’ are used interchangeably to refer to the Russian Soviet Federal Socialist Republic.

⁴The data are from the State Archive of the Russian Federation and the Russian State Archive of the Economy; see the online Data Appendix for details.

population data to calculate birth rates by marital status, marriages by spousal age gap, divorce rates and abortion death rates by five-year age group and region for 1959.⁵

The analysis focuses on individuals age 18 to 44 in 1959, i.e. born in 1915 to 1941. This age range is selected because it is the age during which most decisions regarding marriage and fertility are made; as shown in Figure 1 it also captures the cohorts most affected by the decline in the sex ratio, along with nearby cohorts who were less affected by war deaths. The sex ratio is calculated as the number of men from two years younger to ten years older than women of a given age, divided by the number of women in the same age range. For example, for women age 22 in 1959, the sex ratio is defined as the number of men age 20 to 32 divided by the number of women age 20 to 32; for women age 20-24 in 1959 the sex ratio is calculated as the number of men age 18 to 34 divided by the number of women of the same age.

Descriptive statistics are given in Table 1; Online Appendix Table 2 provides descriptive statistics by five-year age group. The sex ratio is .796 for the population as a whole and varies widely across birth cohorts and regions. A relatively low proportion of women (61.8 percent) were married in 1959, and only 64.8 percent of men. Birth rates are lower for urban than for rural women (81.2 versus 126.5 births per 1,000 women, respectively); the birth rate for married women is 185.5 births per 1,000 married women while the unmarried birth rate is 43.2 births per 1,000 unmarried women. Out of wedlock births accounted for nearly 16 percent of births in Russia in 1959, compared with 5.3 percent of births in the U.S. in 1960. Note that because abortion was legalized in the USSR in 1955 and was a widely-used means of controlling births in the Soviet Union, nonmarital births represent a conscious choice to have a child outside of

⁵Age heaping at ages 30 and 40 is evident in the rural Census data. The sex ratio measure used here, with numerator and denominator spanning the same ages, helps to minimize measurement error due to this problem, providing that men and women are similar in their age misreporting.

marriage. Of marriages registered in 1959, 7.1 percent had a bride and groom with at least an 11-year age gap between them.⁶ The divorce rate was 8.8 per 1,000 married population and is much higher for the urban population (13.1 per 1,000 married population) than for the rural population (3.4 per 1,000). For comparison, the crude divorce rate in Russia was 1.4 per 1,000 population in 1959 versus 2.2 per 1,000 population in the U.S. in 1960. Table 1 also indicates the high number of deaths due to abortion, at .108 per 100 births. Female employment rates were high, with 73.3 percent of women age 16-44 employed in 1959.

Aside from basic indicators such as proportion married and age-specific birth rates, most of these data have never been published; in fact some of the data, such as deaths due to abortion, are marked as “top secret” in the archives and only became available to researchers in the 1990s. As a result, it is difficult to assess the reliability of the data. Where comparable published data are available, the archival data match the published data quite closely. For example, the archival data match the proportion married for men and women age 18 to 44 exactly; the age-specific birth rates are also close (see Online Appendix Table 2). The quality of the underlying Census and vital statistics data is more difficult to assess. The 1959 Census data are considered to be relatively high quality (Anderson et al. 1994), although age heaping is evident in the age distribution data. Registration of vital events was nearly complete in Russia by 1959, but nonmarital births were almost certainly under-registered due to the stigma of having an unnamed father on a child’s birth certificate. Demographers estimate that 5 percent of births were not registered in the USSR in this period, with rural births less accurately recorded than urban births (Anderson & Silver, 1986). Deaths due to abortion are likely under-reported as well.

⁶Since age at marriage is recorded in 5-year age groups, the spousal age gap is difficult to calculate. Here ‘large age gap’ is assigned if, for example, a bride in the 20-24 age group marries a groom in the 35-39 age group or any age group older than that.

III. Empirical strategy and main regression results

The regressions are estimated as fixed effects regressions which stack the data by single year of birth, or 5-year age group, and region:

$$Y_{cj} = \beta_1 R_{cj} + \beta_2 \ln(\text{pop}_{cj}) + \lambda_c + \mu_j + \varepsilon_{cj} \quad (1)$$

where R_{cj} is the sex ratio for birth cohort c in region j ; λ_c are birth cohort fixed effects; μ_j are region fixed effects, and Y_{cj} is an outcome such as the proportion married, the birth rate, divorce rate, and so on. The birth cohort fixed effects are dummy variables for single year of birth or 5-year birth cohort, depending on data availability. As has become standard in this literature, all regressions include the log of the size of the population of the opposite sex which controls for the thickness of the marriage market. The regressions are estimated for women or men age 18-44 in 1959 or a similar age group as dictated by data availability. Because matching in the marriage market may differ between urban and rural areas (the marriage market in urban areas is likely to be ‘thicker’ than in rural areas, and social norms regarding marriage and child bearing may differ between rural and urban areas), the regressions are also estimated for urban and rural areas separately. Standard errors are adjusted for clustering by region, which allows for serial dependence in the errors within regions (Bertrand et al., 2004).

a. Identification

The basic estimating equation (1) includes birth cohort and region fixed effects. The birth cohort effects control for differences across birth cohorts that are common across regions in 1959, such as length of exposure to the 1955 legalization of abortion which varies by year of birth and may have affected past and therefore current fertility decisions. The region fixed effects control for unobserved, unchanging characteristics of Russia’s regions, such as cultural

differences in attitudes toward marriage, that are assumed to be unchanging across birth cohorts but may affect fertility and marital outcomes. In this base specification the coefficient of interest, β_1 , is identified from differences in the sex ratio between single year (or five year) birth cohorts within the same region. The key identifying assumption is that differences in sex ratios across cohorts are exogenous to economic and social conditions that may affect marriage markets and fertility outcomes, and in particular that there are no omitted variables that are correlated with the sex ratio and related to the outcomes in the analysis.

At the national level it is plausible that the change in sex ratios was an exogenous shock: as is well known, the German attack on the USSR in 1941 was a surprise to Soviet leaders and the Soviet population.⁷ Little emigration or immigration was permitted in the early postwar period, so for the country as a whole the sex ratio would have been unaffected by these factors.

However much of the concern regarding the endogeneity of sex ratios and marital outcomes relates to internal migration, i.e. that cross-regional migration within Russia may have occurred in the early postwar period that was correlated with both sex ratios and marriage market conditions in the regions. For example, women faced with unfavorable prospects in the marriage market may have migrated to relatively high sex ratio areas in the Far East, where men earned high wages in the natural resource sector and were attractive marriage prospects. This seems unlikely in the postwar Soviet Union, however. While data are limited, the evidence suggests that cross-regional migration rates were low due to the severe housing shortage in the Soviet Union and the use of residence permits in many cities.⁸ For example, the net urban migration

⁷The German-Soviet Non-Aggression Pact was signed on August 23, 1939. Glantz (1998) describes the unpreparedness of the Red Army for war in 1941.

⁸Residence permits (“*propiska*”) were implemented in 1932 to control population movement. To obtain a *propiska* for a city required confirmed employment in that city; however a *propiska* was

rate for Russia was 24.3 per 1,000 population in 1950 and 16.6 in 1959; the net rural migration rate was even lower (Goskomstat of Russia, 1998). Moreover, cross-regional migration is only weakly correlated with the regional population losses during the war: the correlation between the change in population by region between 1939 and 1951 – a proxy for the physical and population losses suffered in each region – and the urban in-migration rate in 1960 is .08.

It is also striking that across the vast expanse of the Soviet Union, there were virtually no regions that were spared the human loss of the war: every USSR republic, from the Baltics to Central Asia, suffered large losses of life and highly unbalanced sex ratios (see Online Appendix Table 1). Even if migration were free and housing were readily available, few, if any, regions in Russia or the Soviet Union in the postwar aftermath would be attractive to women looking to improve their outcomes on the marriage market. It is more likely that women ‘migrated’ across age groups within their region than migrated across regions to find eligible partners. This response to male scarcity in the marriage market is tested below.

Aside from endogenous internal migration, a second threat to internal validity is from regional economic losses: the regions with the largest population losses likely suffered the greatest economic losses through the physical devastation that resulted from battles in the occupied territories. The economic losses could affect marriage markets indirectly through the impact on the employment prospects of men and women and by further exacerbating the housing shortage. The ethnic structure of the population was altered as well, primarily due to the destruction of the Jewish population (Acemoglu et al., 2011). The full set of region and birth

necessary to be offered a job – a classic Soviet “catch 22” situation. Only urban residents had the right to hold an internal passport; rural residents obtained this right in 1974 and had no right to move even within their oblast of residence. The internal migration that did occur in the USSR in this period was state-driven and mainly comprised employment-led resettlements or forced resettlements of ethnic groups to remote regions of the USSR (Ivakhnyuk, 2009).

cohort dummies in the regressions will absorb much of these effects, and it is unlikely that the impact of the economic losses varied by year of birth in a way that would be correlated with the change in the sex ratio within regions. The loss of the Jewish population affected both men and women so is unlikely to be strongly correlated with the sex ratio. The regressions are also estimated including region-specific linear cohort trends which control for gradually evolving regional characteristics of cohorts. While it would be useful to test the robustness of the results by using war deaths as an instrument for the sex ratio, as in Abramitzky et al. (2011) and Bethmann & Kvasnicka (2013), there are no publicly available data on war deaths by region or age group, and war deaths cannot be reliably estimated due to the 20-year gap between Census enumerations in 1939 and 1959.

b. Results

The first set of regressions uses the share of women or men currently married as the dependent variable; results are shown in Table 2a. Each cell in the table reports the coefficient on *Sex Ratio* from a separate regression. Regressions are estimated without weights (columns 1 and 3) and weighting by population (columns 2 and 4), and including region-specific cohort trends (columns 3 and 4). For women, a lower sex ratio is associated with a lower proportion married; this effect is statistically significant for the population as a whole and for the urban population, but is statistically insignificant for the rural population in most specifications.⁹ Using the base specification in column (1), the coefficient on the sex ratio for women (.149) indicates

⁹In this and following tables, the standard errors for the rural regressions are often much larger than those for the urban regressions, suggesting greater measurement error in the dependent variable for the rural population. This is consistent with the higher incidence of age heaping and the less accurate vital registration reporting for the rural population discussed above.

that the roughly two-standard deviation decrease in the sex ratio that occurred due to the war is associated with a 5.0 percentage point decrease in the share of married women, or an 8 percent decrease in the average marriage rate for women. Including regional cohort trends in the regression increases the magnitude of the coefficients but does not affect the statistical significance of the estimates. The use of population weights has little effect on the magnitude or statistical significance of the results in most specifications, although the OLS estimates are generally more efficient.¹⁰

The results for men are smaller in magnitude and often statistically insignificant in the base specification (columns 1 and 2) but are positive and statistically significant when region-specific cohort trends are included in the regressions (columns 3 and 4). Because cohort trends almost certainly affect the demographic outcomes examined here and likely vary by region, the preferred specification includes the region-specific cohort trends,¹¹ so I focus on these results for men and in further discussion. The coefficient on *Sex Ratio* for men is .178 (column 3), indicating that a two-standard deviation decrease in the sex ratio is associated with a 6 percentage point decrease in the share of men who were married in 1959, or a 9 percent decrease in the average male marriage rate.

This positive relationship between the sex ratio and the male proportion married contradicts Becker's prediction of a negative relationship between the two, and is also the opposite of Abramitzky et al.'s finding (2011) that greater male scarcity in France after WWI increased the probability of male marriage. The unique institutional environment of the Soviet

¹⁰Solon et al. (2015) discuss the use of weights when the data are group averages. While WLS is often employed to correct for heteroskedasticity in group population data, as discussed in Solon et al. the OLS estimator can yield more precise estimates than WLS in some circumstances.

¹¹F-tests reject the null hypothesis that the region-cohort trends are jointly zero in all regressions.

Union may explain this result, as costly divorce and (financially and legally) costless nonmarital fatherhood for men may have deterred men from making legally binding commitments through marriage. Archival research of court records indicates that men preferred to avoid registered marriage and divorce altogether in this era, or, if already in a registered marriage, felt free to leave their wives for another partner without bothering to file for divorce (Nakachi, 2006, 2008).¹² Rather than strengthening the family as intended, the 1944 Family Code appears to have weakened the family and reduced the likelihood that men marry.

To further investigate the impact of family law on marriage in a low sex ratio environment, Table 2b estimates the same regressions as in Table 2a (column 3) using the 1926 Russian Census. The sex ratio was .868 in this period due to the disproportionate loss of men in earlier wars. Data are in 5-year age groups (except for age 18-19); the same regressions using the 1959 Census are shown in columns (3) and (4) for comparison. The 1920s were characterized by ultra-liberal unilateral divorce laws¹³ and the requirement that fathers provide support to all children, legitimate and illegitimate. Both of these significantly reduced the cost of marriage relative to bachelorhood for men compared with the 1959 regime. As shown in Table 2b, the impact of low sex ratios on the proportion of men and women married in 1926 differs markedly from that of 1959. Focusing on the urban population results (which are less affected by age heaping, especially in the 1926 Census), the standard errors of the estimates are almost identical in the 1926 and 1959 regressions. But the coefficients on *Sex Ratio* are much smaller in magnitude and statistically insignificant for the 1926 regressions. This suggests that the impact

¹²This behavior is consistent with the male-female discrepancy in reported marriage in the 1959 Census, in which about 250,000 more women than men report they are currently married.

¹³In 1926 a wife or husband could unilaterally file for divorce in a local registration office; the spouse would be notified of the divorce three days later by postcard (see Moskoff, 1983).

of male scarcity on marriage probabilities is reduced when the cost of marriage is relatively low (although it is also consistent with greater attenuation bias in the 1926 regressions).

Returning to the 1959 results, a second set of regressions tests the effect of the sex ratio on overall birth rates (Table 3). Focusing on the preferred specification in columns 3 and 4, birth rates are higher for women in cohorts with higher sex ratios; the results are qualitatively similar when birth rates are in logs (columns 5 and 6). This result is as expected since most births took place within marriage; the lower marriage rates for women resulting from low sex ratios would therefore result in lower birth rates. Using the coefficient in column (3), a two-standard deviation decrease in the sex ratio is associated with a decline of 54 births per 1,000 women age 18-44, a quantitatively large effect¹⁴ indicating that the lingering impact of low sex ratios on the likelihood of female marriage and therefore birth rates further exacerbated the population loss of World War II by reducing births in the postwar years.

Table 4 shows the results of regressions of the sex ratio on other demographic outcomes; to conserve space, results are shown including region-specific cohort trends only. Columns 1 and 2 test whether the sex ratio affects marital fertility. The results indicate that a lower sex ratio is associated with fewer marital births; in other words, when men are scarce – and therefore have a stronger bargaining position within marriage – the birth rate among married women is lower. One explanation for this relationship may be differing preferences for children, with men preferring fewer children than women and having greater influence on fertility choices given their stronger bargaining position (no data exist on desired fertility of men and woman in Russia

¹⁴This magnitude implies a decline in the total fertility rate (TFR) of 1.4, a large effect compared with the TFR in Russia of 2.63 in 1959. However the birth rate used here is a period rate which can be biased by changes in the timing of births (Newell, 1994). The results for completed fertility, discussed below, are a more reliable indicator of the effect of low sex ratios on fertility.

for this period, however). Another possible explanation is female labor supply, with married women working more in a low sex ratio environment, leading to lower marital birth rates. However, when a control for the female employment to population ratio is included in the regression, the coefficient on the sex ratio remains positive and statistically significant (discussed below). A third possibility is that in regions and cohorts with lower sex ratios, marriages are more unstable: men are more likely to leave a marriage for another partner but remain legally married to the original spouse, leading married women to reduce their fertility. One can conjecture that the low-sex ratio environment of postwar Russia had additional detrimental effects on wives given their weak bargaining power; for example domestic abuse rates might be higher and, with few outside options, women might rationally decide to remain in abusive relationships. However, data limitations preclude statistical testing of these predictions.¹⁵

Unmarried women also had weak bargaining power in the aftermath of the war; given the relative scarcity of men unmarried women may have felt increased pressure to engage in non-marital sexual relations which, due to the lack of contraceptives in the Soviet Union, may have led to an increase in nonmarital births as well as unwanted pregnancies and abortions. A regression of the nonmarital birth rate on the sex ratio supports this prediction, indicating that a two-standard deviation decrease in the sex ratio is associated with an increase in nonmarital births of 68 per 1,000 unmarried women (column 3), a large increase over the average nonmarital birth rate (43 per 1,000). While this effect is quantitatively large, it is plausible given the state financial support of unmarried mothers during this period (the relationship between the nonmarital birth rate and the sex ratio is sensitive to the use of weights, however). The increase

¹⁵A regression of the female homicide rate on the sex ratio (and controls) shows a negative relationship, as predicted, but the coefficient is statistically insignificant in all specifications.

in nonmarital births in Russia contrasts with the effects of war mortality in Bavaria, which indicate a rise in the out of wedlock birth ratio largely driven by declining overall fertility rather than increased nonmarital births (Bethmann & Krasickova 2013). State transfer payments for unwed mothers had declined significantly in postwar Germany, in contrast to Russia, underscoring that differences in the institutional environment likely affect the impact of low sex ratios on nonmarital fertility. Columns (5) and (6) of Table 4 indicate that the share of nonmarital births is negatively related to the sex ratio for urban women, and the effect is economically significant. The coefficient on the sex ratio (column 5) indicates that a two-standard deviation decrease in the sex ratio is predicted to increase the proportion of out of wedlock births by 2.6 percentage points, or 19 percent of the share of nonmarital births (14.0 percent) in urban areas. The effect of the sex ratio in rural areas is statistically insignificant, possibly due to greater measurement error in the reporting of rural nonmarital births.

Urban-rural differences are also evident in the impact of the sex ratio on the spousal age gap (columns 7 and 8) and on the divorce rate (columns 9 and 10). In urban marriage markets, the decline in the sex ratio is associated with a significant increase in the proportion of marriages with a large age gap between bride and groom, and with an increase in the divorce rate (the latter is significant for the urban population in the weighted regression only). These results are consistent with women marrying men outside of the typical age range as a means of adjusting to relative male scarcity. The higher divorce rate could reflect a lower quality of matches in the marriage market, perhaps due to a decrease in assortative matching (Abramitzky et al., 2011). A higher divorce rate in a low sex ratio environment is also consistent with ‘serial polygamy,’ in which better outside options for men lead men to marry, divorce and remarry at higher rates than would be the case in a high sex ratio environment (Becker, 1974, 1981).

The last regressions in Table 4 regress the death rate from abortion on the sex ratio. The death rate from abortion is calculated as the number of deaths due to abortion in each age group divided by the number of births in each age group (to adjust for differing fertility rates); the number of women obtaining abortions is unavailable for 1959. The coefficients on *Sex Ratio* for the whole population and the urban population (columns 11 and 12) are insignificant with large standard errors; the sign for the rural population in both columns is negative, as expected, but is statistically insignificant. As in the divorce data, the abortion data are in 10-year age groups for ages 30-39 and 40-49, which eliminates important sources of variation in the sex ratio. This reduced variation increases the imprecision of the estimates and is compounded by the probable underreporting of abortion deaths. If the true relationship between sex ratios and abortion deaths in rural areas is negative, as suggested by these results, this likely reflects higher abortion rates for women facing lower sex ratios, rather than worse medical care resulting in higher death rates, as there is little reason to think the quality of medical care varies by age group within regions.

To summarize, the results indicate that women and men who faced low sex ratios were less likely to marry and more likely to divorce. Overall birth rates and marital birth rates were lower as well, while nonmarital births increased. The share of marriages with a large age gap between bride and groom increased, and the evidence is suggestive that rural abortions may have increased as well.

IV. Robustness checks and additional evidence

a. Robustness checks

Table 5a presents regressions which test the sensitivity of the results to the inclusion of controls for large region interacted with birth cohort fixed effects. Table 5b examines whether

differences in female labor supply or health are a confounding factor for some of the results.¹⁶

Table 5a replaces the region-specific cohort trends with an interaction between large region dummies and year of birth for each of the 11 large regions in Russia. These large region $\times$ year of birth dummies absorb large region-specific shocks which also vary by age group. While most of the results are similar to those in Tables 2, 3 and 4, one notable difference is that the impact of the sex ratio on the male proportion married (column 2) becomes statistically insignificant for the ‘all’ and urban populations, and is negative and statistically significant for the rural population – similar to the results for men in Table 2a without the region-specific cohort trends – indicating that male marriage probabilities are particularly sensitive to the inclusion or exclusion of these cohort trends.

As argued previously, an omitted variable that would bias the results would have to vary by birth year in a way similar to that of the sex ratio. One variable that might fit this criteria is female labor supply: due to the shortage of male workers during the war, perhaps women were pulled into the labor force (and remained in it) in a way that mimicked the sex ratio.¹⁷ While this seems unlikely – it would require a large difference in the change in female employment for women age 25-29 versus women age 35-39, for example – it can be tested using female employment data from the 1959 Census. Table 5b produces the results of regressions which

¹⁶Online Appendix Table 3 shows results using different definitions of the sex ratio. For most results the statistical significance of the sex ratio coefficient is unchanged, although the coefficient is smaller in magnitude the more narrowly the sex ratio is defined.

¹⁷Goldin & Olivetti (2013) show that female labor force participation in the U.S. increased for more educated women due to World War II. The female labor force participation rate in the USSR was 47.3 percent in 1939 and 48.1 percent in 1950, suggesting a minor impact of World War II on female labor force participation (Ofer & Vinokur, 1985).

include the female employment to population ratio as a control.¹⁸ For most outcomes the magnitude and statistical significance of the *Sex Ratio* coefficient is similar to those of the original specification, indicating that changing female labor supply does not explain the results.

Another possible omitted variable is health status: women in regions and cohorts most affected by the destruction of the war may have suffered greater health shocks through exposure to hunger, trauma and disease; poor health status could in turn reduce one's attractiveness on the marriage market and reduce fecundity.¹⁹ The lasting health impact of exposure to hunger and trauma from World War II in Europe is clearly evident (Kesternich et al., 2014), suggesting that variations in health status by age and region caused by war exposure could explain some of the results. The best available proxy for health status is age-specific death rates (calculated from deaths by 5-year age group in the archival data) which should reflect any differing health effects of exposure to war by region and cohort. Including the female age-specific death rate as a control in the regressions, the sign on this variable is generally as expected (e.g., negative in the birth rate regressions) but is usually statistically insignificant. As shown in Table 5b, the coefficient on *Sex Ratio* is little changed by the inclusion of this proxy for health status. When both female employment and age-specific death rates are included in the regressions, the results are nearly unchanged from those in Table 5b (results shown in Online Appendix Table 4).

¹⁸Note that female employment is an outcome variable, i.e. is predicted to be related to the sex ratio (e.g. Angrist, 2002) and technically should not be included in the regressions as a control variable. When the female employment to population ratio is regressed on the sex ratio and controls, the coefficient on *Sex Ratio* is statistically insignificant.

¹⁹I thank a referee for suggesting this possible source of omitted variable bias.

b. Additional evidence: Belarus and Ukraine

The previous results provide insight into the impact of unbalanced sex ratios on marital status, birth rates, and other outcomes at one point in time. This section uses 1999 and 2001 Census data for Belarus and Ukraine, respectively, to provide retrospective information on possible longer-term consequences of male scarcity, such as completed fertility, childlessness, and the probability of ever marrying.²⁰ Ukraine and Belarus are demographically similar to Russia: for women born in 1915-1941, average completed fertility was 2.30 children in Belarus, 1.97 in Ukraine, and 2.04 in Russia (Online Appendix Table 5; Scherbov & Van Vianen, 2001); the proportion of these women who never married was .052, .046, and .043 in Belarus, Ukraine, and Russia, respectively. In addition, the impact of World War II on the male population was similar in these republics: the sex ratio for the cohort most affected by the war, women age 35-39 in 1959, was .627 in Russia, .631 in Ukraine, and .648 in Belarus (Online Appendix Table 1). Given these similarities, the Census results for Ukraine and Belarus may provide a reasonable indication of the long-term consequences of male scarcity for Russian women. One should note, however, that the results for these outcomes are more susceptible to endogenous migration than the 1959 outcomes. Moreover, healthier women are likely over-represented in the Census, as the respondents comprise women age 57 to 86. If healthier women were more likely to marry and were more fertile, the results may understate the effects of male loss on these outcomes.

Descriptive statistics for the combined Censuses indicate completed fertility of 2.02 children on average for this cohort, with 4.7 percent never marrying and 10.4 percent remaining childless (Table 6). The regressions use the same specification and sex ratio definition as previously; the microdata are collapsed by region to be comparable to the region-level analysis of

²⁰Data are from IPUMS International (<http://international.ipums.org/international>).

the Russian data. All three outcomes – completed fertility and the probability of marriage and childlessness – have the expected signs and are statistically significant at the 5 percent level or less (Table 7). A two-standard deviation decrease in the sex ratio is associated with a decrease in number of children ever born of .15 to .20 child, depending on how the sex ratio is calculated, or a 7.5 to 10 percent decrease in completed fertility. The probability of being childless increased by .031 points or a nearly 30 percent increase in the average rate. The probability of never marrying increased by .034 points with a two-standard deviation decrease in sex ratio, a large effect given the .047 share of women who never married in this cohort. In sum, the results confirm that the effects of the large loss of male life were long-lasting, leading to decreased fertility, lower probability of ever marrying and higher probability of childlessness. The magnitude of the effect on completed fertility is relatively small in light of the scale of male losses, perhaps due to the pronatalist policy that promoted out of wedlock births.²¹

V. Conclusion

World War II exacted a devastating toll on the Soviet population. Tens of millions of people died, mostly men, leaving behind a cohort of women who survived but faced highly unfavorable conditions in the marriage market and within marriage. The results presented in this paper suggest that the effects of the war impacted these women's lives for decades, leading to lower rates of marriage and fertility, and higher rates of out of wedlock births and divorce in urban areas, than would have been the case absent the war. These effects were likely exacerbated

²¹Regressions using the Russian Longitudinal Monitoring Survey test whether low sex ratios increased the number of abortions a woman had in her lifetime. The relationship between the sex ratio and abortions is negative, as predicted, but the sample size is small and the results are sensitive to specification (see Online Appendix Tables 7 and 8).

by an institutional environment that encouraged nonmarital births and increased the cost of binding commitments through marriage, particularly for men. While the 1944 Family Code appears to have increased marital instability and reduced registered marriage for both men and women, it nevertheless may have benefitted women by reducing the stigma of unmarried motherhood and creating a financial and social environment in which unmarried women had a realistic option of having a child.

The shock to sex ratios in the Soviet Union due to World War II was among the largest experienced by any country in the twentieth century. In this sense, the effect on Russian women and men was unique and arguably not relevant to other countries or time periods. Yet highly unbalanced sex ratios characterize many populations – whether due to wars, immigration and emigration, or son preference, among other reasons – and therefore this analysis can shed light on the effects of sex ratio imbalance in other contexts. In particular, the results support the conclusion of a growing body of evidence that sex ratios matter for marital and fertility outcomes, both on the marriage market itself and within marriage. The experience of the Soviet Union, with its backdrop of radically changing divorce and child support laws between the two World Wars, also highlights that the institutional context of the sex ratio imbalance matters for determining both the size and direction of its impact on marriage markets and family formation.

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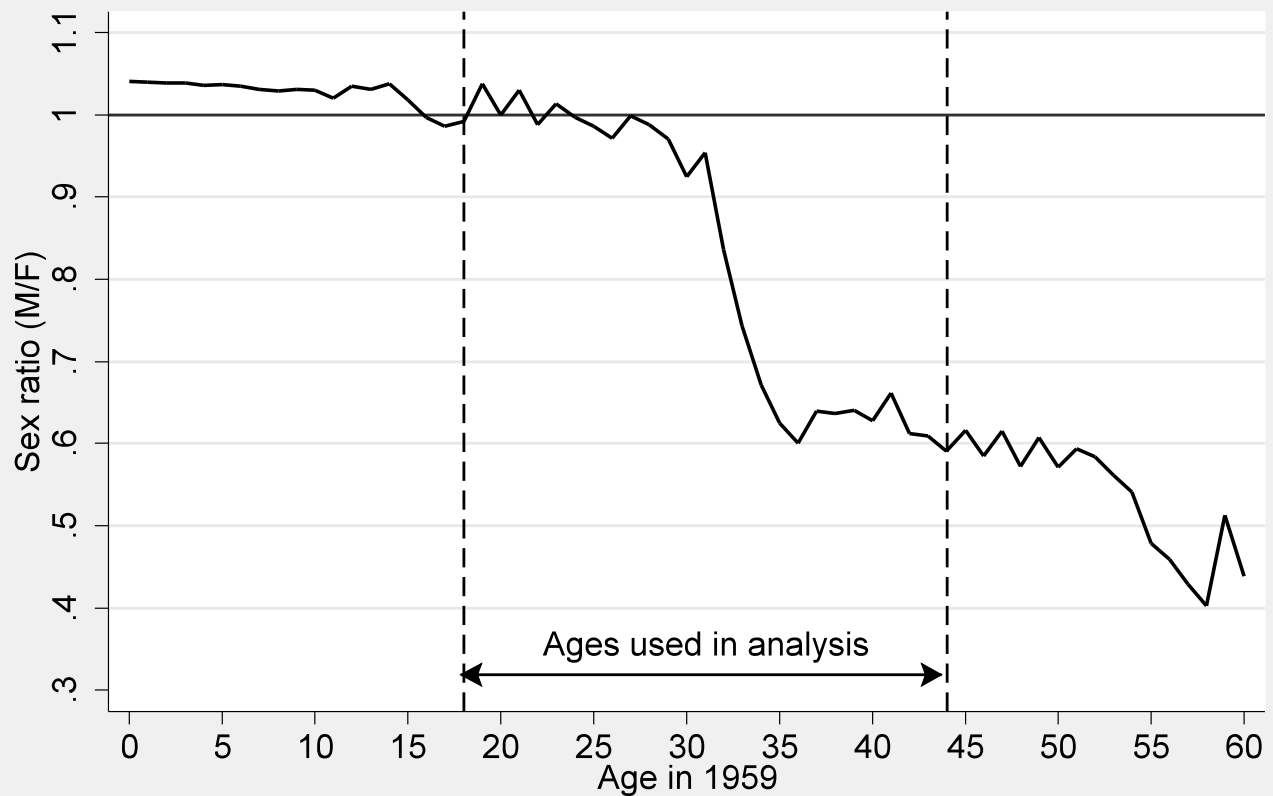
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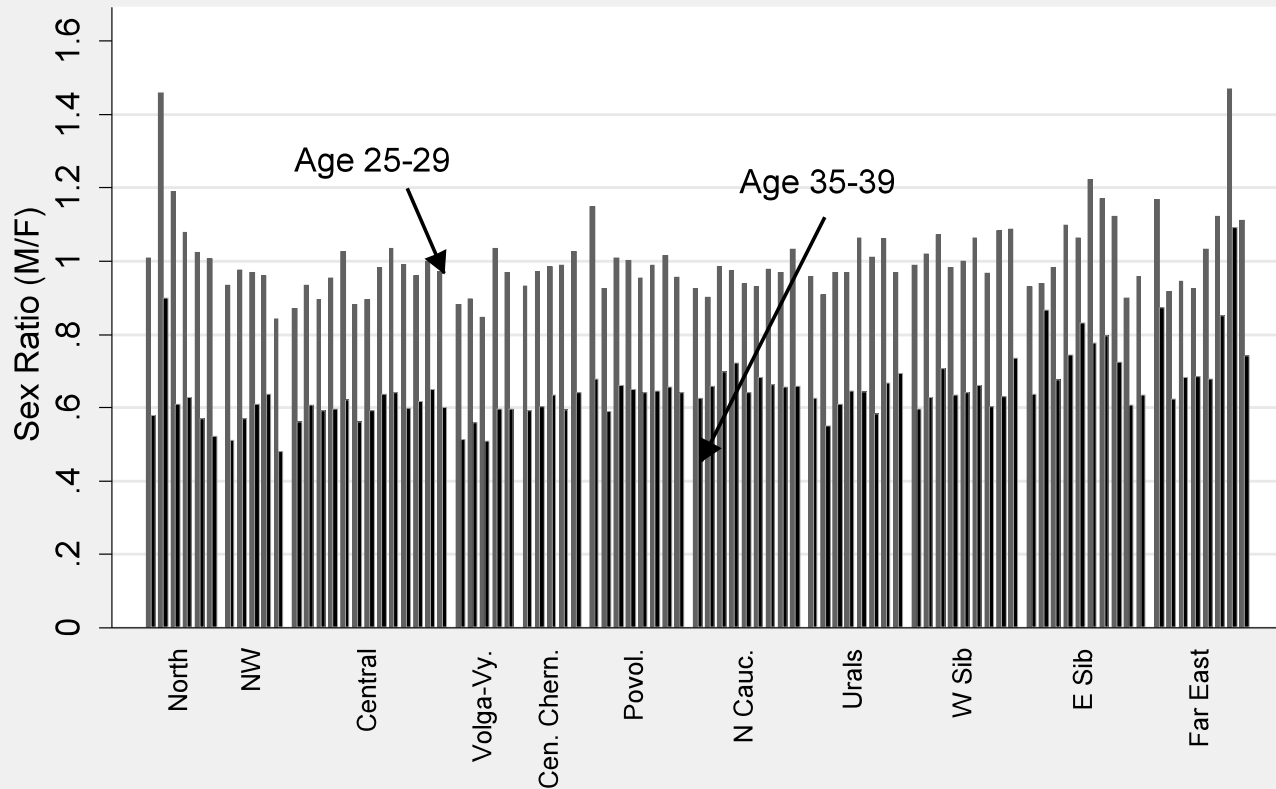
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Fig. 1 Sex Ratio by Single Year of Age, Russia, 1959



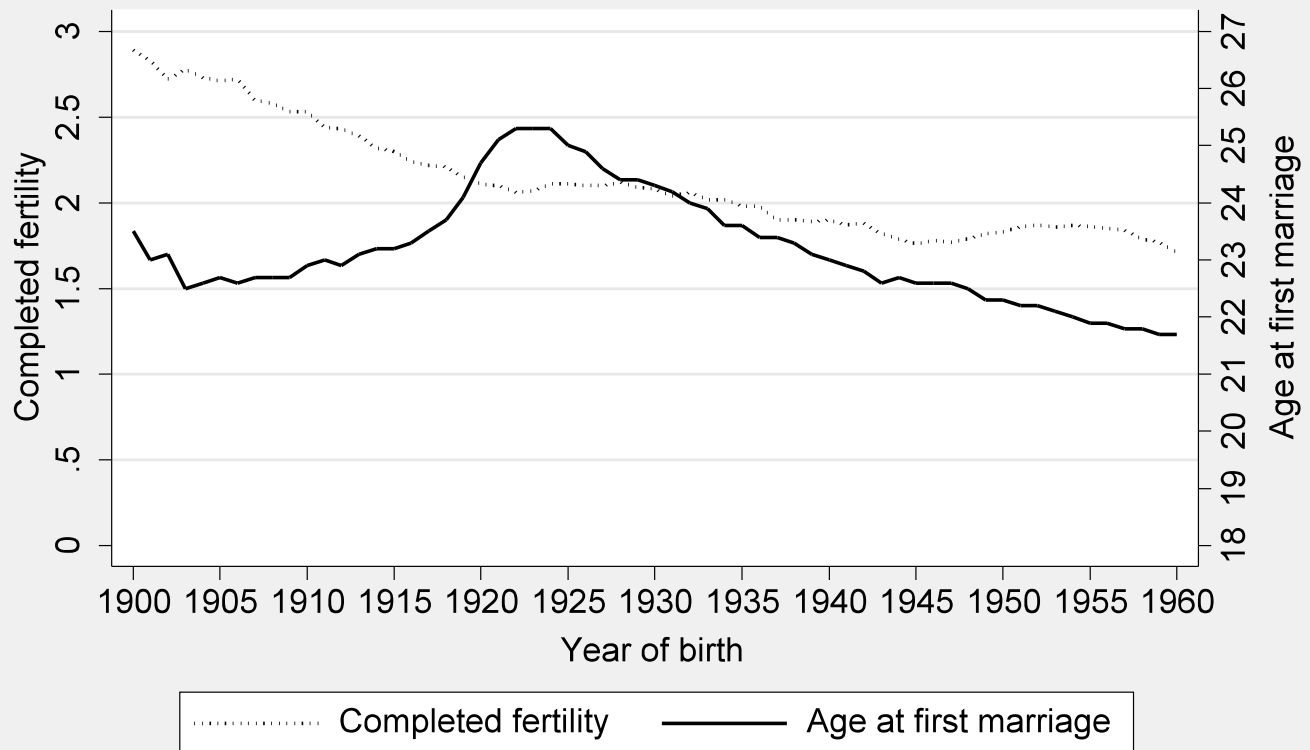
Source: 1959 Soviet Census

Fig. 2 Sex Ratios by Russian Region, Ages 25-29 and 35-39, 1959



Sex ratios here are: (men 25-29)/(women 25-29) and (men35-39)/(women 35-39). Source: 1959 Soviet Census.

Fig. 3 Mean Age at First Marriage and Completed Fertility by Year of Birth, Russian Women



Source: Scherbov and Van Vianen 2001.

Table 1. Summary Statistics, 1959 Russian Census and Vital Statistics
(weighted by population)

	<u>Mean</u>	<u>SD</u>	<u>Min.</u>	<u>Max.</u>	<u>N</u>
<u>AGE 18 - 44, YEAR OF BIRTH 1915-1941; DATA BY SINGLE YEAR OF AGE:</u>					
Sex ratio, all pop.	.796	.168	.447	1.950	2025
Urban pop.	.810	.149	.470	1.995	2025
Rural pop.	.777	.199	.408	1.761	1971
Proportion married:					
Women, all pop.	.618	.201	.034	.893	2025
Urban	.620	.218	.027	.899	2025
Rural	.617	.181	.048	.918	1971
Men, all pop.	.648	.353	.011	.984	2025
Urban	.641	.352	.006	.991	2025
Rural	.658	.355	.013	.987	1971
Births per 1,000 women at each age:					
All	100.6	60.9	.680	305.7	2025
Urban	81.2	52.6	0	313.5	2025
Rural	126.5	71.3	2.8	381.2	1971
<u>AGE 16- 44, YEAR OF BIRTH 1915-1943; DATA BY 5-YEAR AGE GROUP (16-19, 20-24, 25-29, 30-34, 35-39, AND 40-44:</u>					
Sex ratio, all pop.	.809	.162	.459	1.700	437
Urban pop.	.819	.142	.488	1.720	437
Rural pop.	.796	.195	.419	1.661	425
Marital birth rate (births per 1,000 married women):					
All	185.5	129.1	4.1	471.8	437
Urban	171.3	137.9	4.1	506.9	437
Rural	206.2	121.2	13.4	560.7	425
Nonmarital birth rate (nonmarital births per 1,000 unmarried women):					
All	43.2	36.4	0.9	257.5	437
Urban	31.5	27.7	0.9	144.9	437
Rural	58.9	49.2	0.2	340.2	425
Share of out of wedlock births:					
All	.156	.065	.046	.609	437
Urban	.140	.055	.047	.425	437
Rural	.170	.076	.002	.795	425
Proportion of marriages with 11+ year age gap between bride and groom (ages 16-49): ^a					
All	.071	.071	0.0	.355	525
Urban	.068	.068	0.0	.305	525
Rural	.076	.078	0.0	.460	511

Divorce rate (divorces per 1,000 married population) (ages 18-49): ^b					
All	8.8	5.9	0.0	30.4	375
Urban	13.1	7.5	0.0	46.4	375
Rural	3.4	3.2	0.0	20.9	365
Deaths due to abortion per 100 births (ages 16-49): ^c					
All	.108	.114	0.0	.917	365
Urban	.121	.200	0.0	2.37	360
Rural	.104	.122	0.0	.778	345
Female employment to population ratio:					
All	.733	.088	.363	.905	450
Urban	.710	.107	.311	.920	450
Rural	.764	.083	.414	.950	438
Female age-specific death rate (deaths per 1,000 population):					
All	1.7	0.7	0.5	9.4	437
Urban	1.5	0.7	0.3	12.4	437
Rural	1.9	0.8	0.0	16.7	425

^aData are for age groups 16-19, 20-24, 25-29, 30-34, 35-39, 40-44, and 45-49.

^bData are for age groups 18-19, 20-24, 25-29, 30-39, and 40-49.

^cData are for age groups 16-19, 20-24, 25-29, 30-39, and 40-49.

**Table 2a. Stacked regressions by single year of birth, 1959 Russian Census
Year of birth 1915-1941 (age 18-44 in 1959)**

Dependent variable	Women			
	Proportion married			
	OLS (1)	WLS (2)	OLS (3)	WLS (4)
All population (N=2,025)	.149*** (.027)	.150*** (.045)	.201*** (.040)	.187*** (.055)
Urban population (N=2,025)	.111*** (.021)	.104** (.040)	.229*** (.027)	.213*** (.032)
Rural population (N=1,971)	.136*** (.045)	.093 (.066)	.096 (.082)	.118 (.108)
	Men			
All population (N=2,025)	-.026 (.034)	.028 (.031)	.178*** (.049)	.185*** (.058)
Urban population (N=2,025)	.018 (.025)	.076** (.033)	.155*** (.048)	.167*** (.051)
Rural population (N=1,971)	-.075** (.032)	-.048* (.027)	.192*** (.064)	.153** (.063)
Oblast and birth cohort FE	Yes	Yes	Yes	Yes
Oblast-specific linear cohort trend	No	No	Yes	Yes

Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age, divided by the number of women in the same age range. WLS: weighted by population. Regressions control for log(male pop.) (for women) or log(female pop.) (for men) and for % urban ('all population' regressions). Standard errors corrected for clustering within oblasts. ***p < 0.01, **p < 0.05, and *p < 0.1.

**Table 2b. Stacked OLS regressions by 5-year age group,
1926 and 1959 Russian Censuses, ages 18-44**

Dependent variable	Proportion married			
	1926 Census		1959 Census	
	Women	Men	Women	Men
	(1)	(2)	(3)	(4)
All population	.022 (.101)	.105 (.139)	.223*** (.051)	.189*** (.077)
Urban population	.043 (.040)	.031 (.071)	.227*** (.039)	.168** (.072)
Rural population	.040 (.189)	.093 (.176)	.144 (.130)	.174* (.105)
Oblast and birth cohort FE	Yes	Yes	Yes	Yes
Oblast-specific linear cohort trend	Yes	Yes	Yes	Yes

Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age range, divided by the number of women in the same age range. Regressions control for log(male pop.) (for women) or log(female pop.) (for men) and for % urban ('all population' regressions). Standard errors corrected for clustering within oblasts. N=660 for 1926 Census regressions for 'all' population; N=593 for urban; N=623 for rural. N=450 for 1959 Census regressions for 'all' and 'urban' populations'; N=438 for rural population. Source of 1926 data: Tsentral'noe statisticheskoe upravlenie SSSR (Moscow, 1928-1935). ***p < 0.01, **p < 0.05, and *p < 0.1.

**Table 3. Stacked regressions by single year of birth, 1959 Russian Census and Vital Statistics Data
Year of birth 1915-1941 (age 18-44 in 1959)**

Dependent variable	Births per 1000 women:				Log(birth rate):	
	OLS (1)	WLS (2)	OLS (3)	WLS (4)	OLS (5)	WLS (6)
All population (N=2,025)	22.3* (12.5)	30.5 (22.6)	160.7*** (28.8)	167.1*** (43.3)	1.47*** (.211)	.940 (.691)
Urban population (N=2,025)	5.3 (8.6)	11.8 (19.7)	97.0*** (14.0)	93.4*** (16.7)	1.66*** (.224)	1.15** (.558)
Rural population (N=1,971)	30.9 (20.7)	14.2 (37.8)	190.0*** (34.2)	139.5 (103.5)	1.53*** (.266)	1.39*** (.532)
Oblast and birth cohort FE	Yes	Yes	Yes	Yes	Yes	Yes
Oblast-specific linear cohort trend	No	No	Yes	Yes	Yes	Yes

Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age, divided by the number of women in the same age range. WLS: weighted by population. Regressions control for log(male pop.) and % urban population ('all population' regressions). Standard errors corrected for clustering within oblasts. ***p < 0.01, **p < 0.05, and *p < 0.1.

**Table 4. Stacked regressions by 5-year age group, 1959 Russian Census and Vital Statistics Data
Year of birth 1915-1943 (age 16-44 in 1959)**

Dep. variable	Marital birth rate		Nonmarital birth rate		Share of out of wedlock births		Prop. of marriages with 11+ year age gap		Divorce rate		Deaths due to abortion per 100 births	
	OLS (1)	WLS (2)	OLS (3)	WLS (4)	OLS (5)	WLS (6)	OLS (7)	WLS (8)	OLS (9)	WLS (10)	OLS (11)	WLS (12)
All pop.	212.6** (88.9)	235.1* (121.6)	-208.7*** (47.4)	-129.1 (117.6)	-.064 (.059)	-.144 (.089)	-.160*** (.044)	-.122** (.050)	-9.4* (5.2)	-21.1*** (8.1)	.005 (.341)	.336 (.400)
Urban pop.	96.8** (42.4)	83.6** (41.6)	-170.7*** (35.2)	-96.9 (83.8)	-.090* (.050)	-.144** (.070)	-.157*** (.030)	-.180*** (.036)	-9.1 (7.5)	-18.7*** (6.1)	.099 (.554)	.447 (.637)
Rural pop.	202.3** (100.6)	149.1 (210.5)	-222.4** (106.6)	-334.3** (150.8)	.011 (.044)	.034 (.095)	-.002 (.114)	.138* (.083)	.537 (.422)	.624 (.456)	-.067 (.048)	-.047 (.031)

All regressions include oblast and birth cohort fixed effects, and oblast-specific linear cohort trends. Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age range, divided by the number of women in the same age range. WLS: weighted by population. Regressions control for log(male pop.) or log(pop.) (divorce rate); ‘all population’ regressions control for % urban population. Standard errors corrected for clustering within oblasts. See Table 1 for sample sizes and age groups for each dependent variable. ***p < 0.01, **p < 0.05, and *p < 0.1.

Table 5a. Robustness Checks

Dependent variable	Data by 1-year age group:			Data by 5-year age group:				
	Proportion married		Births per 1000 women	Marital birth rate	Nonmarital birth rate	Share of out of wedlock births	Prop. of marriages with 11+ year age gap	Divorce rate
	Women	Men						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
All pop.	.154*** (.047)	-.030 (.055)	76.0*** (29.0)	180.3** (73.9)	-40.1** (20.0)	-.164 (.148)	-.118*** (.044)	-11.2*** (3.9)
Urban pop.	.110*** (.037)	.012 (.043)	25.5* (14.7)	7.5 (49.8)	-44.8** (18.0)	-.018 (.073)	-.122*** (.043)	-6.5 (5.1)
Rural pop.	.142*** (.039)	-.087* (.047)	121.2*** (33.4)	266.4*** (69.6)	-4.4 (31.1)	-.268** (.134)	.002 (.053)	.542* (.285)
Large region × birth cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

All regressions include oblast and birth cohort fixed effects. Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age, divided by the number of women in the same age range. Regressions control for log(male pop.) or log(female pop.) and % urban population (for ‘all population’ regressions). Standard errors corrected for clustering within oblasts. See Table 1 for sample sizes and age groups. ***p < 0.01, **p < 0.05, and *p < 0.1.

Table 5b. Robustness Checks (data by 5-year age group)

Dependent variable	Marital birth rate		Nonmarital birth rate		Share of out of wedlock births		Divorce rate		Deaths due to abortion per 100 births	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)
All population	215.6** (93.0)	214.9** (90.8)	-205.9*** (48.7)	-157.3*** (52.2)	-.063 (.059)	-.092 (.062)	-9.2* (5.2)	-10.7** (5.3)	.016 (.324)	-.042 (.389)
Urban population	113.2*** (41.8)	99.3** (40.6)	-151.1*** (36.5)	-167.7*** (34.9)	-.080 (.058)	-.093* (.050)	-6.9 (7.3)	-8.4 (7.6)	-.056 (.493)	.049 (.556)
Rural population	168.8* (93.5)	227.4** (114.3)	-213.4* (108.7)	-127.2 (99.9)	.025 (.045)	-.007 (.061)	.52 (.44)	.57 (.43)	-.069 (.050)	-.067 (.047)
Control for female employment to pop. ratio	Yes	No	Yes	No	Yes	No	Yes	No	Yes	No
Control for female age-specific death rate	No	Yes	No	Yes	No	Yes	No	Yes	No	Yes

All regressions include oblast and birth cohort fixed effects, and oblast-specific linear cohort trends. Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age range, divided by the number of women in the same age range. Regressions control for log(male pop.) and % urban population (for ‘all population’ regressions). See Table 1 for sample sizes and age groups. Standard errors corrected for clustering within oblasts. ***p < 0.01, **p < 0.05, and *p < 0.1.

Table 6. Summary statistics for Belarus 1999 and Ukraine 2001 Censuses, year of birth 1915-1941 (age 57-86 in 1999/2001)*

	Mean	SD	Min	Max
<u>Belarus 1999 and Ukraine 2001 combined:</u>				
Age	69.2	6.83	57	86
Sex ratio	.793	.141	.469	1.13
Number of children ever born	2.02	.271	1.43	3.52
No children (=1)	.104	.031	.029	.230
Never married	.047	.025	.000	.146
Proportion urban	.529	.239	.088	1.00
Number of obs. in each cell	1304	936	30	5033

*Data are collapsed by region (oblast) of residence and single year of birth. N=918 region × birth year cells (N=758,558 underlying observations). *Sex ratio* is the number of men from -2 to +10 years older than women of a given age in each region, divided by the number of women in the same age range (using 1959 Census data). Data are weighted by the number of observations in each cell. See Online Appendix Table 5 for descriptive statistics for Belarus and Ukraine separately.

**Table 7. Stacked regressions by single year of birth,
Belarus 1999 Census and Ukraine 2001 Census
Year of birth 1915-1941 (age 57-86 in 1999/2001)**

	Dependent variable:		
	Number of children ever born	No children	Never married
Using current region of residence to calculate sex ratio:			
Belarus and Ukraine (n=918)	.536*** (.180)	-.111** (.049)	-.120** (.047)
Using birthplace to calculate sex ratio:			
Belarus and Ukraine (n=1,511)	.724*** (.243)	-.214*** (.055)	-.161** (.061)
Oblast and birth cohort FE	Yes	Yes	Yes
Oblast-specific linear cohort trend	Yes	Yes	Yes

Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age in each region, divided by the number of women in the same age range (using 1959 Census data). Regressions are estimated on Census data collapsed by region (oblast) and year of birth; the underlying number of observations is N=758,558. Regressions are weighted by the number of observations in each cell. Regressions control for log(male pop.) and % urban population. Standard errors corrected for clustering within current region (oblast) of residence or birthplace (oblast or USSR republic). See Online Appendix Table 6 for results for Belarus and Ukraine separately. ***p < 0.01, **p < 0.05, and *p < 0.1.

FOR ONLINE PUBLICATION

Online Appendix Table 1. Population and Sex Ratios by Region and Republic

Region	Population, thous.			1959 Sex Ratio:		
	1939	1951	Change	25-29	35-39	Diff.
USSR republics:						
RSFSR (Russia)	108377	102945	-5432	.983	.627	.356
Ukraine	40469	37223	-3246	.927	.631	.296
Belarus	8912	7781	-1131	.871	.648	.223
Moldova	2452	2392	-60	.857	.779	.078
Estonia	1052	1104	52	.922	.650	.272
Latvia	1885	1954	69	.927	.626	.301
Lithuania	2880	2561	-319	.882	.745	.137
Armenia	1282	1360	78	.949	.735	.214
Azerbaijan	3205	2933	-272	.992	.709	.283
Georgia	3540	3560	20	.908	.655	.253
Kazakhstan	6082	6813	731	1.048	.684	.364
Kyrgyzstan	1458	1764	306	.951	.690	.261
Tajikistan	1485	1554	69	.913	.773	.140
Turkmenistan	1252	1225	27	.963	.697	.266
Uzbekistan	6347	6434	-87	.930	.697	.233
RSFSR oblasts and republics:						
North:						
Republic of Karelia	469	482	13	1.009	.579	.430
Republic of Komi	319	459	140	1.460	.900	.559
Arkhangelskaya oblast	1109	1014	-95	1.191	.610	.581
Vologodskaya oblast	1599	1228	-371	1.024	.569	.455
Murmanskaya oblast	291	337	46	1.008	.521	.487
Northwest:						
St. Petersburg (Leningrad)	3385	2899	-486	.934	.511	.423
Leningradskaya oblast	1294	1000	-294	.975	.569	.406
Novgorodskaya oblast	1152	737	-415	.968	.610	.359
Pskovskaya oblast	1550	1043	-506	.961	.638	.323
Central:						
Bryanskaya oblast	1802	1527	-275	.872	.560	.312
Vladimirskaya oblast	1340	1330	-10	.935	.606	.329
Ivanovskaya oblast	1388	1278	-110	.895	.592	.303
Kaluzhskaya oblast	1178	891	-287	.954	.596	.359
Kostromskaya oblast	1075	923	-152	1.026	.623	.404
City of Moscow	4542	5347	805	.883	.561	.322
Moskovskaya oblast	4255	4131	-124	.895	.591	.303
Orlovskaya oblast	1286	926	-360	.983	.638	.344
Ryazanskaya oblast	1925	1437	-488	1.036	.641	.394
Smolenskaya oblast	1984	1220	-764	.991	.598	.393
Tverskaya oblast	2489	1891	-598	.960	.618	.342
Tulskaya oblast	1729	1734	5	1.000	.651	.348
Yaroslavskaya oblast	1602	1364	-238	.971	.600	.371

Online Appendix Table 1, continued

Volga-Vyatskii:

Mari-el Republic	581	575	-6	.881	.512	.369
Rep. of Mordovia	1185	983	-202	.898	.559	.339
Chuvashskaya Rep.	1078	1026	-52	.846	.509	.337
Kirovskaya oblast	2334	1916	-418	1.035	.597	.439
Nizhegorodskaya oblast	3520	3337	-183	.968	.595	.373

Central Chernozem:

Belgorodskaya oblast	1440	1327	-113	.932	.592	.339
Voronezhskaya oblast	2709	2196	-513	.971	.603	.368
Kurskaya oblast	1773	1418	-355	.985	.635	.351
Lipetskaya oblast	1353	1174	-179	.990	.594	.396
Tambovskaya oblast	1878	1521	-357	1.026	.641	.385

Povolzhsky:

Rep. of Kalmykiya	179	123	-56	1.149	.679	.470
Rep. of Tatarstan	2914	2686	-228	.927	.589	.338
Astrakhanskaya oblast	683	567	-116	1.010	.661	.349
Volgogradskaya oblast	1775	1444	-331	1.002	.650	.352
Penzenskaya oblast	1651	1453	-198	.954	.641	.313
Samarskaya oblast	1646	1809	163	.988	.646	.342
Saratovskaya oblast	2273	1957	-316	1.014	.657	.357
Ulyanovskaya oblast	1183	1108	-75	.956	.643	.313

North Caucasus:

Rep. of Adygeya	247	na	na	.926	.626	.300
Rep. of Dagestan	1023	836	-187	.901	.659	.242
Kabardino-Balk. Rep.	350	344	-6	.985	.698	.287
Karachaevo-Cherk. Rep.	246	na	na	.974	.723	.251
North Ossetia	408	383	-25	.938	.641	.297
Chechnya-Ingushetia	727	459	-268	.929	.683	.256
Krasnodarskii Krai	3172	3338	166	.977	.663	.314
Stavropolskii Krai	1759	1540	-219	.969	.656	.313
Rostovskaya oblast	2893	2756	-137	1.032	.659	.373

Urals:

Rep. of Bashkortostan	3158	2773	-385	.958	.626	.332
Udmurtskaya Rep.	1223	1181	-42	.908	.550	.358
Kurganskaya oblast	976	886	-90	.970	.608	.361
Orenburgskaya oblast	1672	1642	-30	.970	.646	.324
Permskaya oblast	2086	2493	407	1.063	.644	.419
Sverdlovskaya oblast	2610	3268	658	1.061	.667	.394
Chelyabinskaya oblast	1727	2313	586	.970	.694	.276

West Siberia:

Altai Republic	162	na	na	.989	.596	.393
Altaiskii Krai	2388	2409	21	1.019	.630	.390
Kemerovskaya oblast	1654	2181	527	1.071	.708	.363
Novosibirskaya oblast	1862	2060	198	.983	.635	.348

Online Appendix Table 1, continued

Omskaya oblast	1390	1427	37	.999	.642	.357
Tomskaya oblast	643	671	28	1.063	.661	.402
Tyumenskaya oblast	991	1000	9	.968	.603	.364
East Siberia:						
Rep. of Buryatia	546	569	23	.930	.637	.293
Tuva Republic	na	130	na	.938	.867	.071
Rep. of Khakasiya	275	na	na	.983	.677	.306
Krasnoyarskii Krai	1960	2121	161	1.097	.746	.352
Irkutskaya oblast	1303	1428	125	1.171	.797	.374
Chitinskaya oblast	963	819	-144	.900	.608	.293
Far East:						
Sakha Republic	414	377	-37	1.169	.874	.295
Primorskii Krai	888	1036	148	.917	.625	.291
Khabarovskii Krai	657	791	134	.946	.683	.262
Amurskaya oblast	634	618	-16	.925	.685	.240
Kamchatskaya oblast	109	122	13	1.032	.678	.354
Magadanskaya oblast	173	166	-7	1.470	1.091	.379
Sakhalinskaya oblast	100	535	435	1.112	.743	.369
Kaliningradskaya oblast	na	455	na	.842	.480	.362

Sources: *Itogi vsesoyuznoi perepisi naseleniya RSFSR 1959* g., pp. 24-28 and *Naselenie SSSR 1987*, pp. 8-26.

**Online Appendix Table 2. Means by Two-, Four-, Five-, or Ten-Year Age Group,
RSFSR 1959 Census and Vital Statistics Data**
(weighted by population)

Sex ratio, age:	<u>All</u>	<u>Urban</u>	<u>Rural</u>
16-19	1.002	.969	1.048
18-19	1.002	.971	1.046
20-24	.974	.961	.995
25-29	.854	.859	.848
30-34	.738	.755	.716
35-39	.628	.664	.585
40-44	.610	.667	.556

	Women:			Men:		
Proportion married	<u>All</u>	<u>Urban</u>	<u>Rural</u>	<u>All</u>	<u>Urban</u>	<u>Rural</u>
16-17	.024	.016	.034	.004	.003	.006
16-19	.094	.074	.119	.024	.022	.028
18-19	.143	.111	.187	.038	.034	.044
20-24	.479	.447	.527	.269	.251	.293
25-29	.752	.757	.746	.802	.779	.831
30-34	.768	.781	.750	.921	.913	.933
35-39	.712	.736	.679	.953	.948	.958
40-44	.606	.648	.556	.963	.962	.966

	Archival data:			Official published data* (1959-1960):		
Birth rate per 1,000 women of each age:	<u>All</u>	<u>Urban</u>	<u>Rural</u>	<u>All</u>	<u>Urban</u>	<u>Rural</u>
16-19	32.6	26.2	40.3	27.5	23.5	36.8
18-19	50.5	39.4	65.1	na	na	na
20-24	158.2	135.8	191.2	157.7	135.3	192.5
25-29	151.7	125.7	185.5	154.5	127.8	189.7
30-34	101.0	77.2	132.8	100.2	76.3	132.6
35-39	55.6	36.8	81.1	56.5	37.4	82.4
40-44	19.4	10.1	30.6	17.3	10.1	30.6

	Married women:			Unmarried women:		
Birth rate per 1,000 women of each age:	<u>All</u>	<u>Urban</u>	<u>Rural</u>	<u>All</u>	<u>Urban</u>	<u>Rural</u>
16-19	375.9	402.5	353.8	8.0	5.3	11.6
20-24	325.5	310.0	345.8	33.8	22.9	53.1
25-29	179.5	150.2	218.3	76.4	56.8	102.2
30-34	110.4	83.7	147.2	69.6	52.8	90.7
35-39	62.9	40.3	95.4	39.5	27.5	54.2
40-44	25.9	12.3	44.1	10.3	6.0	14.9

Online Appendix Table 2, continued

Proportion of 1959 marriages with 11+ year
age gap between bride and groom by age
of bride

	<u>All</u>	<u>Urban</u>	<u>Rural</u>
16-19	.018	.020	.016
20-24	.005	.005	.004
25-29	.012	.012	.012
30-34	.050	.049	.051
35-39	.129	.132	.126
40-44	.172	.168	.177
45-49	.168	.153	.187

Divorces per 1,000
married population:

	<u>All</u>	<u>Urban</u>	<u>Rural</u>
18-19	6.7	11.7	2.3
20-24	13.5	20.4	5.1
25-29	11.9	17.6	4.7
30-39	8.3	11.8	3.5
40-49	3.5	5.2	1.2

Deaths due to abortion
per 100 births:

	<u>All</u>	<u>Urban</u>	<u>Rural</u>
Age 16-19	.065	.070	.062
20-24	.036	.035	.038
25-29	.057	.057	.059
30-39	.111	.117	.108
40-49	.208	.260	.184

Female employment
to population ratio:

	<u>All</u>	<u>Urban</u>	<u>Rural</u>
Age 16-19	.620	.537	.725
20-24	.790	.759	.835
25-29	.767	.768	.765
30-34	.738	.733	.745
35-39	.735	.720	.756
40-44	.698	.664	.740

Female age-specific death rate
(per 1,000 pop.)

	<u>All</u>	<u>Urban</u>	<u>Rural</u>	<u>Official data, all pop. (1958-59)*</u>
Age 15-19	1.0	0.8	1.2	1.0
20-24	1.2	1.0	1.5	1.2
25-29	1.4	1.2	1.6	1.4
30-34	1.7	1.5	1.9	1.7
35-39	2.2	2.0	2.4	2.2
40-44	2.9	2.8	3.0	2.9

*Source: *Naselenie Rossii za 100 let (1897-1997) (Population of Russia: 1897-1997)*, pp. 145-47, 150-51.
For births, the official data are for births to mothers age ≤ 15 through age 19 and are therefore not comparable to the age 16-19 category for the archival data.

Online Appendix Table 3. Additional Robustness Checks

All population, unweighted OLS regressions

	Data by 1-year age group:			Data by 5-year age group:				
Dependent variable	Proportion married		Birth rate	Marital birth rate	Nonmarital birth rate	Share of out of wedlock births	Prop. of marriages with 11+ year age gap	Divorce rate
	Women	Men						
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Sex ratio: base specification	.201*** (.040)	.178*** (.049)	160.7*** (28.8)	212.6** (88.9)	-208.7*** (47.4)	-.064 (.059)	-.160*** (.044)	-9.4* (5.3)
Sex ratio: narrow definition	.153*** (.030)	.076** (.031)	108.0*** (21.0)	125.2* (63.9)	-158.8*** (38.0)	-.050 (.051)	-.131*** (.029)	-9.2** (4.1)
Sex ratio: broad definition	.251*** (.077)	.295*** (.079)	246.0*** (38.2)	382.4*** (146.7)	-211.0** (81.8)	-.078 (.087)	-.144 (.105)	-10.8 (7.5)
Oblast and birth cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Oblast-specific linear cohort trend	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age, divided by the number of women in the same age range (base specification). The narrow definition is the number of men from -2 to +5 years older than women of a given age, divided by the number of women in the same age range, while the broad definition is men from -2 to +15 years older than women of a given age, divided by the number of women in the same age range. All regressions control for log(male pop.) (for women) or log(female pop.) (for men) and for % urban ('all population' regressions). Standard errors corrected for clustering within oblasts. N=2025 or N=1971 for regressions on data by one-year age group. Marital and nonmarital birth data are for the 16-19, 20-24, 25-29, 30-34, 35-39 and 40-44 age groups (N=437 for all and urban regressions; N=425 for rural regressions). 11+ year marital age gap data are for the 16-19, 20-24, 25-29, 30-34, 35-39, 40-44 and 45-49 age groups (N=525 for all and urban regressions; N=511 for rural regressions). Divorce data are for the 18-19, 20-24, 25-29, 30-39 and 40-49 age groups (N=375 for all and urban regressions; N=365 for rural regressions). . ***p < 0.01, **p < 0.05, and *p < 0.1.

Online Appendix Table 4. Additional Robustness Checks

	Data by 5-year age group:									
Dependent variable	Marital birth rate		Nonmarital birth rate		Share of out of wedlock births (father undeclared)		Divorce rate.		Deaths due to abortion per 100 births	
	OLS	WLS	OLS	WLS	OLS	WLS	OLS	WLS	OLS	WLS
All population	219.3** (93.4)	209.8* (108.2)	-152.8*** (49.9)	-120.0 (74.0)	-.090 (.063)	-.146 (.096)	-10.6** (5.3)	-19.6*** (5.7)	-.051 (.371)	.285 (.384)
Urban population	114.9*** (41.1)	86.1** (41.3)	-149.0*** (36.9)	-49.4 (62.8)	-.082 (.062)	-.124 (.079)	-6.3 (7.5)	-18.4** (7.1)	-.114 (.492)	.149 (.575)
Rural population	188.6* (111.4)	97.0 (185.8)	-113.8 (93.4)	-244.0* (133.9)	.008 (.061)	.054 (.097)	.53 (.44)	.75 (.53)	-.069 (.049)	-.051 (.034)
Control for female emp. to pop. ratio	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Control for female age-specific death rate	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Oblast and birth cohort FE	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes
Oblast-specific linear cohort trend	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes	Yes

Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age range, divided by the number of women in the same age range. WLS: regression weighted by population. All regressions control for log(male pop.) and % urban (for ‘all population’ regressions). Divorce regressions also control for the male age-specific death rate. Standard errors corrected for clustering within oblasts. Marital birth data are for the 16-19, 20-24, 25-29, 30-34, 35-39 and 40-44 age groups (N=437 for all and urban regressions; N=425 for rural regressions). Divorce data are for the 18-19, 20-24, 25-29, 30-39 and 40-49 age groups (N=375 for all and urban regressions; N=365 for rural regressions). . ***p < 0.01, **p < 0.05, and *p < 0.1.

**Online Appendix Table 5: Summary statistics for Belarus 1999 and Ukraine 2001 Census data
Year of birth 1915-1941 (age 57-86 in 1999/2001)***

	Mean	SD	Min	Max
<u>Belarus 1999:</u>				
Age	67.4	6.74	57	83
Sex ratio	.758	.111	.507	.945
Number of children ever born	2.30	.278	1.65	3.12
No children (=1)	.101	.031	.045	.191
Never married	.052	.022	.006	.123
Proportion urban	.522	.223	.197	1.0
Number of obs. in each cell	810	228	124	1265
<u>Ukraine 2001:</u>				
Age	69.6	6.78	59.9	86
Sex ratio	.801	.145	.469	1.13
Number of children ever born	1.97	.230	1.43	3.52
No children (=1)	.104	.030	.029	.230
Never married	.046	.025	0	.146
Proportion urban	.531	.243	.088	1.0
Number of obs. in each cell	1407	993	30	5033

*Data are collapsed by region (oblast) of residence and single year of birth. N=189 region × birth year cells for Belarus (N=130,881 underlying observations); N=729 region × birth year cells for Ukraine (N=627,677 underlying observations). *Sex ratio* is the number of men from -2 to +10 years older than women of a given age, divided by the number of women in the same age range. Data are weighted by the number of observations in each cell.

**Online Appendix Table 6. Stacked regressions by single year of birth,
Belarus 1999 Census and Ukraine 2001 Census
Year of birth 1915-1941 (age 57-86 in 1999/2001)**

	Dependent variable:		
	Number of children ever born	No children	Never married
	Using current region of residence to calculate sex ratio:		
Belarus (n=189)	.774 (.473)	-.222** (.074)	-.171** (.066)
Ukraine (n=729)	.458** (.189)	-.095 (.059)	-.126** (.056)
	Using birthplace to calculate sex ratio:		
Belarus (n=443)	1.68** (.621)	-.329*** (.111)	-.250** (.097)
Ukraine (n=1,068)	.479* (.248)	-.191*** (.059)	-.157** (.069)
Oblast and birth cohort FE	Yes	Yes	Yes
Oblast-specific linear cohort trend	Yes	Yes	Yes

Each cell shows the coefficient on *Sex Ratio* from a separate regression. *Sex ratio* is the number of men from -2 to +10 years older than women of a given age, divided by the number of women in the same age range. Regressions are estimated on Census data collapsed by region (oblast) and year of birth; the underlying number of observations is N=130,881 for Belarus and N=627,677 for Ukraine. Regressions are weighted by the number of observations in each cell. All regressions control for log(male pop.) and % urban population. Standard errors corrected for clustering within current region (oblast) of residence or birthplace (oblast or USSR republic). . ***p < 0.01, **p < 0.05, and *p < 0.1.

Online Appendix: Russian Longitudinal Monitoring Survey results

Summary statistics for the RLMS variables are given in Table 7. Nearly three-quarters of women report ever having had an abortion, with three abortions on average. These figures are plausible given the high abortion rates in Russia, but the distribution of abortions is implausible with some heaping evident at 10, 15, and 20 abortions (and 37 abortions almost physically impossible). Given these problems, a 0/1 dummy variable for *Ever had an abortion* may produce results less subject to recall or misreporting problems.

The regressions take the same form as previous regressions, with the fertility outcome regressed on the sex ratio, log of the male population, oblast and year of birth fixed effects, and oblast-specific linear cohort trends in some regressions. The results are reported in Table 8. The relationship between total births and the sex ratio is positive, as predicted, but the coefficient is statistically insignificant in all specifications. The coefficient on number of abortions is negative, as predicted, but again the coefficients are statistically insignificant. Using a median regression to minimize the problem of outliers produces similar results, as does a Poisson regression to account for the non-negative count data (results not shown). Column (3) shows the results using *Ever had an abortion* has the dependent variable. The coefficient is negative and statistically significant in the first row, but becomes statistically insignificant when oblast cohort trends are included in the regression. Given the small sample size, the recall problems, and measurement error in the sex ratio it is unsurprising that these coefficients are imprecisely estimated; however the negative sign on *Sex Ratio* in the abortion regressions is at least suggestive that abortions increased in the low sex ratio aftermath of World War II.

Online Appendix Table 7. Descriptive statistics for RLMS data, women age 53-83
(Year of birth 1915-1941)

	Mean	SD	Min	Max
Age	64.0	7.2	53	83
Sex ratio	.821	.174	.461	1.55
Number of children ever born	2.22	1.43	0	14
Ever had an abortion (=1)	.741	.438	0	1
Number of abortions	3.05	3.95	0	37
Median number of abortions	2			
Share of respondents not answering abortion questions	.091			
Number of observations	2,090			

Sex ratio is the number of men from -2 to +10 years older than women of a given age in each region, divided by the number of women in the same age range (using 1959 Census data, based on the current region of residence). Data source: 1994-1998 rounds of the RLMS (www.cpc.unc.edu/projects/rlms-hse).

Online Appendix Table 8. RLMS regressions, women age 53-83
(Coefficient on *Sex Ratio* is shown in each cell)

Dependent variable:	Total number of births (1)	Number of abortions (2)	Ever had an abortion? (1 = yes) (3)
With oblast and birth cohort fixed effects	0.87 (1.09)	-2.50 (3.75)	-.871** (.338)
Adding oblast-specific linear cohort trend	3.30 (2.22)	-1.02 (3.08)	-.590 (.463)
N	1,959	1,880	1,883

Sex ratio is the number of men from -2 to +10 years older than women of a given age in each region, divided by the number of women in the same age range (using 1959 Census data, based on the current region of residence). Regressions also include controls for rural area, log (male population), and total live births (columns 2 and 3). Weighted using survey weights. Standard errors clustered by region in parentheses. ***p < 0.01, **p < 0.05, and *p < 0.1.

Online Data Appendix: Archival data sources

The archival data used in this paper were hand-collected and digitized by the author from two archives in Moscow: the GARF archive (*Gosudarstvennyi arkhiv Rossiiskoi Federatsii* (State Archive of the Russian Federation)) and the RGAE archive (*Rossiiskii gosudarstvennyi arkhiv ekonomiki* (Russian State Archive of the Economy)). The specific location of each data series is as follows (F. = *Fond*; op. = *opis*; d. = *delo*):

Age and sex distribution of the population by marital status, single year of age and RSFSR oblast, 1959: GARF F. 374, op. 40, d. 1, 2, 3, 4 and RGAE F. 1562, op. 336, d. 1531, 1532, 1533, 1534.

Age and sex distribution of the population by marital status, single year of age and oblast for all other USSR republics: RGAE F. 1562, op. 336, d. 1535-1548.

Total births by single year of age and RSFSR oblast, 1959: GARF F. 374, op. 31, d. 4921.

Total births by marital status and five-year age group and RSFSR oblast, 1959: GARF F. 374, op. 31, d. 4923.

Registered marriages in 1959 by age of bride and groom and RSFSR oblast: GARF F. 374, op. 31, d. 4943.

Divorces in 1959 by age and RSFSR oblast: GARF F. 374, op. 31, d. 4942.

Deaths from abortion by 2-, 5- or 10-year age group and RSFSR oblast, 1959: RGAE F. 1562, op. 27, d. 834, 835, 836.

Male and female employment by 5-year age group and RSFSR oblast, 1959: RGAE F. 1562, op. 336, d. 2760, 2775-2825.

Female deaths by five-year age group and RSFSR oblast, 1959: RGAE F. 1562, op. 27, d. 826.